1	The T_{peak} – T_{end} interval as an electrocardiographic risk marker of arrhythmic and
2	mortality outcomes: a systematic review and meta-analysis
3	Short title: Systematic review of $T_{peak} - T_{end}$ for risk stratification
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61 Abstract

Background: The T_{peak} – T_{end} interval, an electrocardiographic marker reflecting transmural
dispersion of repolarization, has been used to predict ventricular tachycardia/fibrillation (VT/VF)
and sudden cardiac death (SCD) in different clinical settings.

65 **Objective**: This systematic review and meta-analysis evaluated the significance of $T_{peak} - T_{end}$ 66 interval in predicting arrhythmic and/or mortality endpoints.

67 Methods: PubMed, Embase, Cochrane Library and CINAHL Plus databases were searched
68 through 30th November 2016.

Results: Of the 854 studies identified initially, 33 observational studies involving 155856 69 patients were included in our meta-analysis. T_{peak} - T_{end} interval prolongation (mean cut-off: 70 103.3 ± 17.4 ms) was a significant predictor of the arrhythmic or mortality outcomes (odds ratio 71 (OR): 1.14, 95% CI: 1.11 to 1.17, p < 0.001). When different end-points were analyzed, the ORs 72 are as follows: VT/VF (1.10, 95% CI: 1.06 to 1.13, p < 0.0001), SCD (1.27, 95% CI 1.17 to 1.39, 73 p < 0.0001), cardiovascular death (1.40, 95% CI 1.19 to 1.64, p < 0.0001), and all-cause 74 75 mortality (4.56, 95% CI 0.62 to 33.68, p < 0.0001). Subgroup analysis for each disease revealed that the risk of VT/VF or death was highest for Brugada syndrome (OR: 5.68, 95% CI: 1.57 to 76 20.53, p < 0.01), followed by hypertension (OR: 1.52, 95% CI: 1.26 to 1.85, p < .0001), heart 77 failure (OR: 1.07, 95% CI: 1.04 to 1.11, p < .0001) and ischemic heart disease (OR: 1.06, 95% 78 CI: 1.02 to 1.10, p = 0.001). In the general population, a prolonged $T_{peak} - T_{end}$ interval also 79 predicted arrhythmic or mortality outcomes (OR: 1.59, 95% CI: 1.21 to 2.09, p < 0.001). 80

81 **Conclusion**: The $T_{peak} - T_{end}$ interval is useful risk stratification tool in different diseases and in 82 the general population.

84 Introduction

Ventricular arrhythmias can take the form of monomorphic or polymorphic ventricular 85 tachycardia (VT) or ventricular fibrillation (VF). Both are life-threatening, potentially 86 culminating in sudden cardiac death (SCD). SCD is a major health problem with a devastating 87 impact on both economic and social issues. The prevalence of SCD is high with up to 60,000 88 deaths in the U.K.¹, 200,000 deaths in the U.S.² and 4 to 5 million deaths worldwide ³, annually. 89 Reliable stratification markers are therefore of paramount importance in identifying high risk 90 patients for SCD. Several electrocardiographic (ECG) markers related to increased risk of 91 arrhythmias and SCD have been proposed ⁴⁻⁶. Traditional ECG markers of ventricular 92 repolarization including the corrected QT (QT_c) interval ⁷ and QT dispersion (QT_D) ⁸ have been 93 used for risk stratification in various clinical settings. Relatively new ECG markers of ventricular 94 repolarization, such as the interval from the peak to the end of the T wave $(T_{peak} - T_{end})^9$, and the 95 $(T_{peak} - T_{end})/OT$ ratio ¹⁰, have been recently proposed to predict ventricular arrhythmic events 96 and SCD¹¹. These ECG markers have been validated in congenital ion channelopathies such as 97 Long QT and Brugada syndromes ¹²⁻¹⁴, myocardial infarction ¹⁵, cardiomyopathies ¹⁶ and other 98 diseases such as pulmonary embolism, hypertension and Chagas disease ^{17, 18}. However, data are 99 controversial regarding the predictive value of these ECG markers ¹⁹⁻²³. The present systematic 100 review and meta-analysis of the current literature aimed to investigate the prognostic significance 101 of T_{peak} – T_{end} interval with respect to arrhythmic and mortality outcomes. 102

103

104 Method

The meta-analysis was performed according to the Preferred Reporting Items for 106 Systematic Reviews and Meta-Analyses statement ²⁴. MEDLINE, Embase, Cochrane library and 107 CINAHL Plus were searched for studies that investigated the relationship between $T_{peak} - T_{end}$ 108 109 interval with arrhythmic or mortality endpoints using the following terms: ["Tpeak – Tend" OR "Tpeak-Tend" OR "Tp - Te" OR "Tp-Te" OR "Tpeak-end" OR "Tp-e" OR "T(peak)-T(end)" 110 OR "T wave peak-to-end" OR "T peak-T end" OR "TPEc" OR "T-peak to T-end" OR "Tpeak-to-111 tend"]. The search period was from the beginning of the databases (1965 for PubMed, 1910 for 112 Embase, 1996 for Cochrane Library, 1937 for CINAHL Plus) through to 30th November 2016, 113 114 with no language restrictions. The following inclusion criteria were applied: i) the design was a case-control, prospective or retrospective observational study in humans, ii) T_{peak} – T_{end} interval 115 durations were determined; iii) endpoint events [appropriate implantable cardioverter-116 117 defibrillator therapy (ICD), ventricular tachycardia/fibrillation (VT/VF), sudden cardiac death (SCD), cardiovascular death (CVD) or all-cause mortality were reported and iv) odds ratios 118 (ORs) or hazard ratios (HRs) and the corresponding 95% confidence intervals (CIs) or data 119 necessary to calculate these were described. 120

The quality assessment of these studies included in our meta-analysis was performed using the Newcastle–Ottawa Quality Assessment Scale (NOS) ²⁵. The point score system evaluated the categories of study participant selection, comparability of the results, and quality of the outcomes. The following characteristics were assessed: a) representativeness of the exposed cohort; b) selection of the non-exposed cohort; c) ascertainment of exposure; d) demonstration that outcome of interest was not present at the start of study; e) comparability of cohorts on the basis of the design or analysis; f) assessment of outcomes; g) follow-up period sufficiently long for outcomes to occur; and h) adequacy of follow-up of cohorts. This scale varied from zero to nine stars, which indicated that studies were graded as poor quality if they met <5 criteria, fair if they met 5 to 7 criteria, and good if they met >8 criteria. The details of the NOS quality assessment are shown in Supplementary Tables 1 and 2.

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133 Data extraction and statistical analysis

134 Data from the different studies were entered in pre-specified spreadsheet in Microsoft 135 Excel. All potentially relevant reports were retrieved as complete manuscripts and assessed for compliance with the inclusion criteria. In this meta-analysis, the extracted data elements 136 137 consisted of: i) publication details: last name of first author, publication year and locations; ii) 138 study design; iii) follow-up duration; iv) definition of T_{peak} - T_{end} interval; v) lead(s) where the T_{peak} - T_{end} interval was measured; vi) endpoint(s); vii) the quality score; and viii) the 139 140 characteristics of the population including sample size, gender, age and number of subjects. Meta-analyses of observational studies are challenging due to differences in study designs and 141 142 inherent biases. This systematic review was conducted in accordance to PRISMA statement ²⁶ and registered with PROSPERO (Review number 52916). Two reviewers (GT and MG) 143 independently reviewed each included study and disagreements were resolved by adjudication 144 145 with input from a third reviewer (TL).

The endpoints of the study were the occurrences of ventricular arrhythmias (VT/VF), SCD, cardiovascular death or all-cause mortality. The definition of these endpoints used in the different studies were analyzed. If more than one mortality endpoint was described, then SCD was preferentially used for analysis, followed by cardiovascular death and all-cause mortality. Multivariate adjusted odds ratios (OR) or hazard ratios (HR) with 95% confidence interval (CI) were extracted and analyzed for each study. When values from multivariate analysis were not available, those from univariate analysis were used. When the latter were not provided, raw data were used to calculate unadjusted risk estimates where possible. Where arrhythmic or mortality outcomes were determined but ORs or HRs were not reported, we contacted the corresponding authors of the studies. HR value in multivariate Cox proportional hazards model was equated as OR. The pooled adjusted risk estimates from each study as the OR values with 95% CI were presented.

The heterogeneity between studies was determined using Cochran's Q, the weighted sum 158 of squared differences between individual study effects and the pooled effect across studies, and 159 the I^2 statistic from the standard chi-square test, which describes the percentage of the variability 160 in effect estimates resulting from heterogeneity, rather than sampling error. $I^2 > 50\%$ was 161 considered to reflect significant statistical heterogeneity. A fixed effects model was used if $I^2 <$ 162 163 50%, otherwise the random-effects model using the inverse variance heterogeneity method was used. To locate the origin of the heterogeneity, sensitivity analysis excluding one study at a time, 164 and subgroup analyses based on different disease conditions and different endpoints were 165 performed. Funnel plots, Begg and Mazumdar rank correlation test and Egger's test were used to 166 assess for possible publication bias. 167

168

169 **Results**

A flow diagram detailing the above search terms with inclusion and exclusion criteria is depicted in Figure 1. A total of 401, 310, 27 and 122 entries were retrieved from PubMed, Embase, Cochrane Library and CINAHL Plus, respectively. Comparing with the entries extracted from the PubMed search, 143, 23 and 116 duplicate entries from the Embase, Cochrane

library and CINAHL Plus searches were found and removed. This yielded 854 publications and
further assessment demonstrated that 30 met the inclusion criteria ^{5, 10, 15, 22, 27-50}. Three groups
provided their data on odds ratio (OR) or hazard ratio (HR), and these studies were also included.
Thus, in the final meta-analysis, 33 studies were included.

A total of 155856 patients were included. Three studies examined the risk in different 178 patient populations (normotensive and hypertensive; dilated cardiomyopathy and ischemic 179 cardiomyopathy; normal intraventricular conduction and intraventricular conduction delay). The 180 $T_{\text{peak}} - T_{\text{end}}$ interval was examined in the following clinical settings: heart failure in eight studies 181 27, 31, 35, 38, 40, 41, 45, 48, ischemic heart disease in eight studies 15, 22, 36, 39, 40, 43, 49, 53, Brugada 182 syndrome in six studies ^{5, 10, 29, 34, 44, 50}, hypertension in two studies ^{30, 51}, pulmonary embolism in 183 one study ³³, Chagas disease in one study ³⁷, intraventricular conduction delay in one study ⁴², 184 dilated cardiomyopathy in one study ⁴⁰ and ischemic cardiomyopathy in one study ⁴⁰. Five 185 studies addressed the prognostic significance of $T_{peak} - T_{end}$ interval in the general population ^{28,} 186 ^{30, 32, 42, 46}. The baseline characteristics of these studies are listed in Table 1. Fifteen were 187 prospective studies and 14 were retrospective studies. The mean follow-up duration was 42 ± 48 188 months. 189

In the 33 studies, the total number of patients was 155856 (mean: 4329; range from 23 to 138404). The mean age was 62 ± 11 years old). The subjects were predominantly male (69%). The mean cut-off point for $T_{peak} - T_{end}$ interval was 103.3 ± 17.4 ms (range between 77.4 and 146.4 ms). All studies consistently reported a positive association between increased $T_{peak} - T_{end}$ interval and increased risk of VT/VF or SCD (17 using multivariate analysis and 16 using univariate analysis). The pooled meta-analysis demonstrated that prolonged $T_{peak} - T_{end}$ interval is associated with 1.14 times higher risk of VT/VF or SCD (95% CI: 1.11 to 1.17, p < 0.0001;

Figure 2). The Cochran's O value was greater than the degrees of freedom (432 vs. 34), 197 suggesting the true effect size was different among the various studies. Moreover, I^2 took a value 198 of 92%, suggesting significant heterogeneity was present. Funnel plot plotting standard errors or 199 200 precision against the logarithms of the odds ratio are shown in Figures 3 and 4, respectively. Begg and Mazumdar rank correlation suggested no significant publication bias (Kendal's Tau 201 value 0.15, p > 0.05). Egger's test demonstrated significant asymmetry (intercept 3.5, t-value 8.1; 202 P < 0.0001)⁵². When HR and OR were analyzed separately, the data were as follows: HR = 1.12 203 (95% CI: 1.09 to 1.16, p < 0.0001; Figure A1); OR = 1.23 (95% CI: 1.14 to 1.32, p < 0.0001; 204 205 Figure A2).

206 To locate the origin of the heterogeneity, sensitivity analysis excluding one study at a time, and subgroup analyses based on different disease conditions and endpoints were 207 performed. Sensitivity analysis by the leave-one-out method did not affect the overall odds ratio 208 209 when each study was removed. VT/VF and different mortality measures were subsequently analyzed as different end-points. For spontaneous and inducible VT/VF, the OR was 1.10 (95% 210 CI: 1.06 to 1.13, p < 0.0001) (Figure A3). Exclusion of three studies reporting inducible VT/VF 211 212 did not significantly alter the pooled OR (1.09, 95% CI: 1.06 to 1.13; Figure A4). For mortality endpoints, the ORs were: SCD (1.27, 95% CI 1.17 to 1.39, p < 0.0001; Figure A5), 213 cardiovascular death (1.40, 95% CI 1.19 to 1.64, p < 0.0001; Figure A6), and all-cause mortality 214 (4.56, 95% CI 0.62 to 33.68, p < 0.0001; Figure A7). Subgroup analyses based on diagnosis were 215 216 subsequently performed.

217

218 *Heart failure*

For heart failure, eight studies ^{27, 31, 35, 38, 40, 41, 45, 48} consisting of 1912 patients (range from 219 84 to 572) with a mean age of 64 ± 13 years (72% males) were included. The mean follow-up 220 period was 21 \pm 14 months. The mean cut-off point for $T_{peak} - T_{end}$ interval was 106.3 \pm 8.4 ms. 221 All eight groups consistently reported a positive association between increased T_{peak} - T_{end} 222 interval and increased risk of VT/VF or SCD (7 using multivariate analysis and 1 using 223 univariate analysis). The pooled meta-analysis demonstrated that prolonged $T_{peak} - T_{end}$ interval 224 was associated with approximately 1.07 times the risk of these endpoints (95% CI: 1.04 to 1.11, 225 p < 0.0001; Figure A8). The Cochran's Q value was greater than the degrees of freedom (56 vs. 226 6), which would suggest different true effect size among different studies. I^2 took a value of 227 88%, suggesting most of the observed variance reflects heterogeneity between studies. 228

229

230 Ischemic heart disease

231 For ischemic heart disease, data from eight studies involving 3402 subjects were included in the sub-group analysis $^{15, 22, 36, 39, 40, 43, 49, 53}$. The mean age was 63 ± 12 years old (77% males). 232 The mean follow-up period was 18 ± 12 months. The mean cut-off point for $T_{peak} - T_{end}$ interval 233 was 109.6 \pm 20.4 ms. All eight studies consistently reported a positive association between 234 increased T_{peak} – T_{end} interval and increased risk of VT/VF or SCD (three studies using 235 236 multivariate analysis and five studies using univariate analysis). The pooled meta-analysis demonstrated that prolonged $T_{peak} - T_{end}$ interval is associated with approximately 1.06 times the 237 risk of these endpoints (95% CI: 1.02 to 1.10; p < 0.001) (Fig. A9). The Cochran's Q value was 238 239 greater than the degrees of freedom (51 vs. 6), indicating the true effect size were different among different studies. A I^2 value of 89.6% suggested that most of the observed variances 240 reflect differences in true effect sizes. 241

243 Brugada syndrome

244 For Brugada syndrome, six studies involving 583 subjects were included (range from 23 to 325) $^{5, 10, 29, 34, 44, 50}$. The mean age was 46 ± 11 years old and 80% of subjects were male. The 245 mean follow-up period was 50 \pm 8 months. The mean cut-off point for $T_{peak} - T_{end}$ interval was 246 95.8 ± 16.3 ms. All six studies consistently reported a positive association between increased 247 T_{peak} - T_{end} interval and increased risk of VT/VF or SCD (2 using multivariate analysis and 4 248 using univariate analysis). The pooled meta-analysis demonstrated that prolonged $T_{peak} - T_{end}$ 249 interval is associated with approximately 5.68 times the risk of these endpoints (95% CI: 1.57 to 250 20.53, p < 0.001; Fig. A10). The Cochran's Q value was greater than the degrees of freedom (35) 251 vs. 5), indicating that differing true effect sizes among the different studies. An I^2 of 86% 252 253 suggests high heterogeneity.

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255 Hypertension

For hypertension, two studies involving 881 subjects were included (range from 57 to 824) $^{30, 51}$. The mean age was 51 ± 11 years old and 55% of subjects were male. The mean follow-up period was 192 months. The mean cut-off point for T_{peak} – T_{end} interval was 96.7 ± 36.3 ms. Both studies consistently reported a positive association between increased T_{peak} – T_{end} interval and increased risk of VT/VF or SCD in multivariate analysis. The pooled meta-analysis demonstrated that prolonged T_{peak} – T_{end} interval is associated with approximately 1.52 times the risk of these endpoints (95% CI: 1.26 to 1.85, p < 0.01; Fig. A10). The Cochran's Q value was 263 greater than the degrees of freedom (1.1 vs. 1), indicating that differing true effect sizes among 264 the different studies. An I^2 of 6% suggests a low heterogeneity.

265

266 *General population*

For the general population, five studies involving 148215 subjects (mean age 62 ± 11 267 years old, 43% males) were included (ranges from 65 to 138404)^{28, 30, 32, 42, 46}. The mean follow-268 269 up period was 111 ± 55 months. The mean cut-off point for $T_{peak} - T_{end}$ interval was 99.8 ± 27.6 ms. All five studies consistently reported a positive association between increased $T_{peak} - T_{end}$ 270 interval and increased risk of VT/VF or SCD (2 using multivariate analysis and 3 using 271 272 univariate analysis). The pooled meta-analysis demonstrated that prolonged $T_{peak} - T_{end}$ interval is associated with approximately 1.6 times higher risk of reaching these endpoints (95% CI: 1.2 273 274 to 2.1, p < 0.0001; Figure A12). The Cochran's Q value was less than the degrees of freedom (25) vs. 4), indicating that differing true effect sizes among the different studies. An I^2 value of 84.0% 275 276 suggests a high heterogeneity among studies.

277

278 Discussion

The main findings of this study are the following:

i. A prolonged T_{peak} – T_{end} is associated with a 1.14 fold increased risk in VT/VF,
 SCD, cardiovascular death or all-cause mortality when data from all pathological
 conditions were pooled with significant heterogeneity among studies;

ii. Subgroup analyses demonstrated that the risk of VT/VF and/or SCD in Brugada
syndrome was the highest with a 5.6 fold increase compared to 1.52 in
hypertension, 1.07 in heart failure and 1.06 in ischemic heart disease.

286 iii. In the general population, a prolonged $T_{peak} - T_{end}$ interval was also predictive of 287 arrhythmic or mortality outcomes with an OR of 1.59.

The cellular origin of the T-wave has been an area of intense study the previous decades 288 ⁵⁴⁻⁵⁶. The waveform has been attributed to electrophysiological characteristics of ventricular 289 cardiomyocytes located in the different regions of the myocardial wall, such as epicardium, mid-290 myocardium (M) and endocardium 57. T_{peak} – T_{end} is defined as the interval between the peak of 291 the T wave and the end of the T wave, representing the dispersion of repolarization ⁹. Initially, it 292 was hypothesized that the T_{peak} - T_{end} interval reflects the transmural dispersion of repolarization 293 (TDR). Later work found that the end of epicardial repolarization coincided with T_{peak} and end of 294 M-cell repolarization coincided with T_{end} ⁵⁸. Subsequent experiments in pigs demonstrated that 295 T_{peak} coincided with the earliest end of repolarization, whereas T_{end} coincided with the latest end 296 297 of repolarization. In other words, T_{peak} - T_{end} was a measure of global dispersion of repolarization rather than TDR 9, 59-61. T_{peak} - T_{end} is also lead-dependent as the dispersion of 298 repolarization varies with different cardiac regions ⁶². Therefore, for left ventricular diseases, 299 300 measurements from lead V5 and for right ventricular diseases such as Brugada syndrome, 301 measurements from lead V2, have been used for ECG interval analysis. In some studies, T_{peak} -302 T_{end} were calculated from mean values of all 12 leads. Although the mechanism of the T wave generation remains controversial, as to whether it represents global or transmural dispersion of 303 repolarization, a prolonged T_{peak} - T_{end} interval has been associated with an increased incidence 304 of ventricular tachyarrhythmias ^{5, 10, 15, 22, 27-50, 63}. Increased spatial dispersion of repolarization 305

can produce unidirectional block, which predisposes to circus-type or spiral reentry ^{60, 64-66}. 306 Moreover, this may reflect loss of the action potential dome in the epicardial region compared to 307 the endocardial region. This is expected to increase the risk of phase 2 reentry ^{67, 68}. Several ECG 308 parameters, such as QT interval, QT dispersion and T-wave alternans (TWA) are associated with 309 T_{peak} - T_{end}. The occurrence of TWA is expected to increase the spatial dispersion of 310 repolarization. Indeed, microvolt TWAs have been associated with the duration of $T_{peak} - T_{end}^4$. 311 The mechanism of TWA generation is multi-factorial but has traditionally been described by the 312 restitution hypothesis ⁵. The TWA magnitude is likely a function of the heterogeneity in Ca²⁺ 313 314 alternans which can drive APD alternans. Conversely, a steep spatial gradient of repolarization can convert spatially concordant alternans to spatially discordant alternans. 315

The prognostic significance of T_{peak} – T_{end} interval has been investigated in various 316 clinical settings. A prolonged T_{peak} - T_{end} interval has been associated with increased 317 arrhythmogenicity in Long QT syndrome (LQTS)1 and LQTS2 at baseline ⁶⁹. Exercise is known 318 to trigger ventricular arrhythmias in LQTS1 but not LQTS2. Greater increases in T_{peak} - T_{end} 319 interval were observed in LQTS1, suggesting that it could be a useful risk marker for 320 arrhythmogenesis in this LQTS subtype. An accentuation of the T_{peak} - T_{end} interval has been 321 associated an increased propensity to develop Torsades de Pointes (TdP) in subjects with LQTS1 322 ¹². The $T_{peak} - T_{end}$ interval is also increased in Short QT syndrome (12). There are limited data 323 regarding the utility of T_{peak} - T_{end} interval in Brugada syndrome ^{10, 13, 14, 50}. A prolonged T_{peak} -324 T_{end} interval has been associated with arrhythmic events in Brugada syndrome ⁵⁰, which is 325 326 consistent with pre-clinical data that TDR is involved in arrhythmogenesis in Brugada syndrome $^{70\text{-}73}\text{.}$ Previous studies have underscored the prognostic significance of T_{peak} – T_{end} interval in 327 328 subjects with structural heart disease including hypertrophic cardiomyopathy and myocardial

infarction. The Copenhagen study found an inverted U relationship between $T_{peak} - T_{end}$ interval and the risk of all-cause and cardiovascular mortality, atrial fibrillation and heart failure ³². However, the abitily of $T_{peak} - T_{end}$ interval to predict prognosis or arrhythmic events has not always been successful ^{19-21, 23}. Moreover, shortenings of this interval also predicted worsened survival rates ⁷⁴.

As shown in our meta-analysis, a prolonged $T_{peak} - T_{end}$ interval displays the highest predictive ability for arrhythmic events in Brugada syndrome compared to other clinical conditions.

In Brugada syndrome, both the depolarization and repolarization hypotheses have been proposed to explain the abnormal electrophysiological findings ^{71, 75}. This would lend weight towards abnormal repolarization being a significant contributor to arrhythmic substrate. On the contrary, in heart failure patients, there is only a small, albeit significant, increase in arrhythmic risk. This possibly suggests that increased dispersion of repolarization plays a moderated role in ventricular arrhythmogenesis, and other factors such as abnormal action potential restitution ⁷⁶ or conduction abnormalities may be more important ⁷⁷.

It should be noted that the results are not dramatic. Based on this meta-analysis we would 344 advocate that a different cut-off value should be considered for each cardiac pathology which 345 346 should also be considered alongside other known factors known to associate with cardiac risk such as such as QT interval, QT dispersion or T wave alternans ⁷⁸. Increased dispersion of 347 repolarization, which is reflected by the prolonged $T_{peak} - T_{end}$ intervals, is only one mechanism 348 349 by which re-entrant mechanism is generated. Indeed, in Mines' seminal work on circus-type reentry, his proposal included three criteria: the presence of unidirectional conduction block, a 350 351 distinct pathway along which the cardiac excitation can propagate, and interruption of the circuit

will terminate the re-entrant activity ⁷⁹. Prolonged $T_{peak} - T_{end}$ interval will increase the 352 likelihood of generating unidirectional conduction block, but other factors, such as slowed 353 conduction and increased dispersion of conduction are also important but not reflected in the 354 T_{peak} – T_{end} interval. A recent meta-analysis showed that another measure of repolarization, the 355 QT interval, predicted mortality⁸⁰. The results were more dramatic, reporting a 24% increase in 356 the risk of SCD with every 50 millisecond increase in the QT interval. 357

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Cut-off points for different conditions

Of the different study populations, the degree of $T_{peak} - T_{end}$ prolongation for significant 360 361 elevations in arrhythmic risk for the general population is the greatest with a cut-off point of 362 113.6 ms. For some diseased states, the cut-off value is much lower. Thus, for Brugada syndrome and heart failure, the cut-off values for T_{peak} - T_{end} duration were 95.8 ms and 106.3 363 364 ms, respectively. Interestingly, the cut-off for ischemic heart disease patients was not significantly different form that of the general population, with a value of 109.6 ms. Whilst the 365 T_{peak} - T_{end} could provide additional information for risk stratification, at the moment it should 366 not be used on its own in estimating arrhythmia risk. However, it could provide incremental 367 information regarding risk stratification in more complex patients and when the risk estimation 368 369 based on conventional parameters might be difficult to calculate.

370

371 Limitations

This systematic review with meta-analysis has several potential limitations. Firstly, 372 hazard ratios were equated as odds ratios. It has been suggested that when event rates or 373

probabilities are low, hazard ratios and odds ratios can be equated ⁸¹. Nonetheless, we have 374 performed additional analysis by pooling HRs and ORs separately. Secondly, a significant 375 heterogeneity among studies was noted. Sensitivity analysis removing one study at a time did not 376 377 alter the pooled odds ratio. Therefore, in the overall meta-analysis, the heterogeneity is likely derived from the distinct patient populations with different diseases. Thirdly, publication bias in 378 meta-analyses is frequently examined by checking for asymmetry in a funnel plot. In our case 379 380 there was significant asymmetry, which may suggest some bias. However, it is known that effect estimates such as odd ratios used in this meta-analysis correlate with standard errors, and can 381 produce asymmetry in a funnel plot. Fourthly, some studies included in our studies are 382 retrospective studies, which may have more recall bias. Finally, although the overall number of 383 patients included in this meta-analysis is large, for certain conditions such as Brugada syndrome 384 a small number of patients (500 patients) were included potentially affecting or masking the true 385 effect. Finally, our systematic review only included articles published in PubMed, Embase, 386 Cochrane and CINAHL. It therefore might have missed articles that were not indexed in these 387 search engines. 388

389

390 Tables

Table 1. Characteristics of the 33 studies included in the meta-analysis.

392

393 Appendices

Figure A1. Forest plot demonstrating the hazard ratios for studies examining the relationship between $T_{peak} - T_{end}$ and arrhythmic or mortality outcomes. **Figure A2.** Forest plot demonstrating the odds ratios for studies examining the relationship between $T_{peak} - T_{end}$ and arrhythmic or mortality outcomes.

Figure A3. Forest plot demonstrating the odds ratios for studies reporting inducible orspontaneous VT/VF outcomes.

400 Figure A4. Forest plot demonstrating the odds ratios for studies reporting spontaneous VT/VF401 outcomes.

402 **Figure A5.** Forest plot demonstrating the odds ratios for studies reporting sudden cardiac death.

Figure A6. Forest plot demonstrating the odds ratios for studies reporting cardiovascular death.

Figure A7. Forest plot demonstrating the odds ratios for studies reporting all-cause mortality.

Figure A8. Forest plot demonstrating the association between $T_{peak} - T_{end}$ and arrhythmic or mortality outcomes in patients with heart failure.

Figure A9. Forest plot demonstrating the association between $T_{peak} - T_{end}$ and arrhythmic or mortality outcomes in patients with ischemic heart disease.

Figure A10. Forest plot demonstrating the association between $T_{peak} - T_{end}$ and arrhythmic or mortality outcomes in patients with Brugada syndrome.

Figure A11. Forest plot demonstrating the association between $T_{peak} - T_{end}$ and arrhythmic or mortality outcomes in patients with hypertension.

Figure A12. Forest plot demonstrating the association between $T_{peak} - T_{end}$ and arrhythmic or mortality events in the general population.

415 **Supplementary Table 1**. NOS risk of bias scale for case-control studies.

416 **Supplementary Table 2**. NOS risk of bias scale for cohort studies.

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Figure 1. Flow diagram of the study selection process.

Figure 2. Forest plot demonstrating the association between $T_{peak} - T_{end}$ and arrhythmic or mortality outcomes in patient populations with different clinical conditions.

Figure 3. Funnel plot of standard errors against logarithms of odds ratios.

Figure 4. Funnel plot of precision measure against logarithms of odds ratios.