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Antipyretic Therapy in Critically Ill Septic Patients: A Systematic Review and Meta-Analysis

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Objective: This meta-analysis aimed to examine the impact of antipyretic therapy on mortality in critically ill septic adults.

Data Sources: Literature searches were implemented in Ovid Medline, Embase, Scopus, Cumulative Index of Nursing and Allied Health Literature, Cochrane Central Register of Controlled

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Trials, NHS Economic Evaluation Database, and ClinicalTrials.gov through February 2016.

Study Selection: Inclusion criteria were observational or randomized studies of septic patients, evaluation of antipyretic treatment, mortality reported, and English-language version available. Studies were excluded if they enrolled pediatric patients, patients with neurologic injury, or healthy volunteers. Criteria were applied by two independent reviewers.

Data Extraction: Two reviewers independently extracted data and evaluated methodologic quality. Outcomes included mortality, frequency of shock reversal, acquisition of nosocomial infections, and changes in body temperature, heart rate, and minute ventilation. Randomized and observational studies were analyzed separately.

Data Synthesis: Eight randomized studies (1,507 patients) and eight observational studies (17,432 patients) were analyzed. Antipyretic therapy did not reduce 28-day/hospital mortality in the randomized studies (relative risk, 0.93; 95% CI, 0.77–1.13; $I^2 = 0.0\%$) or observational studies (odds ratio, 0.90; 95% CI, 0.54–1.51; $I^2 = 76.1\%$). Shock reversal (relative risk, 1.13; 95% CI, 0.68–1.90; $I^2 = 51.6\%$) and acquisition of nosocomial infections (relative risk, 1.13; 95% CI, 0.61–2.09; $I^2 = 61.0\%$) were also unchanged. Antipyretic therapy decreased body temperature (mean difference, -0.38°C ; 95% CI, -0.63 to -0.13 ; $I^2 = 84.0\%$), but not heart rate or minute ventilation.

Conclusions: Antipyretic treatment does not significantly improve 28-day/hospital mortality in adult patients with sepsis. (*Crit Care Med* 2017; 45:806–813)

Key Words: acetaminophen; antipyretics; fever; mortality; sepsis

Over 1 million patients are hospitalized with sepsis annually in the United States, and sepsis is the leading cause of death in critically ill patients (1). Fever, a common sign of infection, occurs in approximately 40% of critically ill septic patients at some point during their ICU stay (2, 3). It is an extremely complex physiologic response with potentially beneficial and harmful effects in septic patients. Fever boosts several aspects of innate and adaptive immunity,

inhibits microorganism growth, slows viral replication, and augments antibiotic efficacy (4–8). In animal models, artificially raising core body temperature leads to improved survival and lower infectious burden (9, 10). However, fever generation also raises the metabolic rate, increases oxygen consumption, and can adversely affect cardiac function (11–13). In septic patients, who are vulnerable to malperfusion and tissue hypoxia, this physiologic expense could be particularly detrimental.

Despite the potential benefits of fever in patients with sepsis, treatment with antipyretic therapies is common in the ICU. In a recent international survey of ICU practitioners in 23 countries, greater than 80% of respondents reported controlling fever in critically ill patients most or all of the time (14). Data supporting this practice, however, remain inconclusive because of limited sample sizes and lack of reproducibility of study results. In fact, some studies have suggested that antipyresis in critically ill septic patients may be harmful (15–17). The majority of prior meta-analyses of the effect of antipyretic therapy in the critically ill have not focused on septic patients (18–20). Because antipyretic therapy may impact infected and noninfected patients differently (16), conclusions from these studies are difficult to interpret. Furthermore, methodologic limitations in previous evaluations of antipyretic therapy in sepsis render the question of optimal fever management in this population unclear (21).

The objective of this meta-analysis was to evaluate the effect of antipyretic therapy on mortality in critically ill septic patients. Secondary aims included assessing the impact of fever control on the acquisition of nosocomial infections, shock reversal, and physiologic variables such as body temperature, heart rate, and minute ventilation. The primary hypothesis was that antipyretic therapy would not improve mortality in septic patients.

MATERIALS AND METHODS

This meta-analysis was conducted and reported in accordance to the Preferred Reporting Items for Systematic Reviews and Meta-Analysis and Meta-Analysis of Observational Studies in Epidemiology guidelines (**Supplemental Digital Content 1**, <http://links.lww.com/CCM/C421>) (22, 23). The study protocol (**Supplemental Digital Content 2**, <http://links.lww.com/CCM/C422>) was developed prior to initiation of the search strategy and has been registered on PROSPERO (registration number: CRD42016037622). Ethical approval from the human research protection office was not required.

Literature Search and Study Selection

Published literature was electronically searched by two medical librarians (A.C.H., S.A.F.) for the concepts of sepsis, fever, antipyretics, and physical cooling in adults. These strategies were implemented in Ovid Medline, Embase, Scopus, Cumulative Index of Nursing and Allied Health Literature, Cochrane Database of Systematic Reviews, Cochrane Central Register of Controlled Trials, NHS Economic Evaluation Database, and ClinicalTrials.gov between January 1946 and February 2016. Full search strategies are provided in **Supplemental Digital Content 3** (<http://links.lww.com/CCM/C423>).

Two authors (A.M.D., E.T.M.) independently screened titles and abstracts for potentially eligible studies. These included observational or randomized studies evaluating mortality in septic patients treated with and without antipyretic therapy. The full list of inclusion and exclusion criteria is available in **Supplemental Digital Content 4** (<http://links.lww.com/CCM/C424>). Studies of antipyretic therapy that included both septic and nonseptic patients were included if mortality data were provided for the subgroup of septic patients. If these data were not reported, authors were contacted via electronic mail to request it. The authors (A.M.D., E.T.M.) also reviewed bibliographies of included articles and performed a hand search of critical care–related journals to identify additional studies. Abstracts from critical care–related meetings (full list provided in the protocol) from 2008 to 2015 were searched to identify any unpublished literature. Any article identified by either screener as being potentially eligible was reviewed in full.

Following the initial screening, full articles were independently reviewed by two authors (A.M.D., B.M.F.) with application of the same inclusion and exclusion criteria. Disagreements regarding study inclusion were resolved by consensus. Studies excluded after full-text review are listed in **Supplemental Table 1** (**Supplemental Digital Content 5**, <http://links.lww.com/CCM/C425>).

Data Extraction

Data on study characteristics, patient characteristics, study interventions, and outcomes were independently extracted from each study by two study members (A.M.D., E.A.A.) using standardized forms created in an online data management system (24). A full list of variables collected is provided in **Supplemental Digital Content 2** (<http://links.lww.com/CCM/C422>). Primary data reported solely in graphical form were extracted using an online plot data extraction tool (25). When necessary, authors were contacted to provide missing data. Following extraction, data were compared and disagreements were resolved by consensus.

Outcomes

The primary outcome was 28-day mortality. Studies reporting hospital mortality were pooled with those reporting 28-day mortality. Secondary outcomes included “early” mortality (defined as mortality on or prior to day 14 after enrollment or within the ICU), frequency of acquisition of nosocomial infections, frequency of shock reversal, and mean changes in body temperature, heart rate, and minute ventilation with antipyretic treatment. A priori, the decision was made to analyze 28-day and 14-day mortality separately based on observations of different mortality rates for these different follow-up periods (26, 27). For randomized trials, postintervention physiologic values were pooled for meta-analysis rather than the pre- to postchange in those values because no study provided measures of dispersion for the pre- to postintervention changes. This was considered to be a valid approach based on the assumption that in randomized trials, the differences in mean final values are similar to the differences in changes of these values (28).

Quality Assessment

For randomized trials, study quality was assessed independently by two reviewers (A.M.D., E.A.A.) using the Cochrane Collaboration Risk of Bias Tool with standardized criteria for evaluating bias in seven domains (29). Quality of observational studies was evaluated with the Newcastle-Ottawa Scale (NOS), a 9-point scale assessing bias in the areas of patient selection, comparability, exposure, and outcome (30). Disagreements were resolved by a third reviewer (B.M.F.). A priori, it was decided that randomized studies with a high or unclear risk of bias in less than two domains or observational studies with an NOS score greater than 7 would be considered to be high quality.

Data Analysis

Observational and randomized studies were analyzed separately, as recommended by expert opinion (31), using STATA/

IC 14.1 (StataCorp, College Station, TX). For categorical outcomes, a relative risk (RR) with 95% CI (for randomized studies) or odds ratio (OR) with 95% CI (for observational studies) was calculated for each study. Data were combined using the DerSimonian and Laird (32) random effects model and plotted as forest plots. Higgins I^2 tests were used to assess heterogeneity. A random effects model was used even if no heterogeneity was observed due to limitations of statistical tests for heterogeneity. For observational studies, adjusted ORs, if available, were preferentially used in the meta-analysis. For studies evaluating multiple methods of antipyresis, the overall OR for any type of antipyresis was used in the meta-analysis. However, if an overall OR was not reported, ORs for each method of antipyresis were included separately. For continuous outcomes, weighted mean differences were calculated using a random effects model for continuous outcomes.

For continuous data reported as median and interquartile range, mean and SD were estimated using previously published methods (33). A p value less than 0.05 was considered statistically significant.

Publication bias was assessed using funnel plots and Egger test. Extended funnel plots were created to graphically display the effect size and SE combinations needed for an additional randomized trial to change the results of the meta-analysis (34, 35). Simulation methods were used to create a graph demonstrating the power achieved by an additional randomized trial to change the results of the meta-analysis at different sample sizes up to a maximum of 30,000 patients (36, 37).

Stratified analyses were conducted for the primary outcome by the type of intervention, duration of treatment, and primary goal of the study (evaluation of anti-inflammatory treatment or evaluation of fever treatment). Predefined subgroup analyses for the primary outcome were performed for the subset of studies with a low risk of bias and for the subset of patients with fever and septic shock.

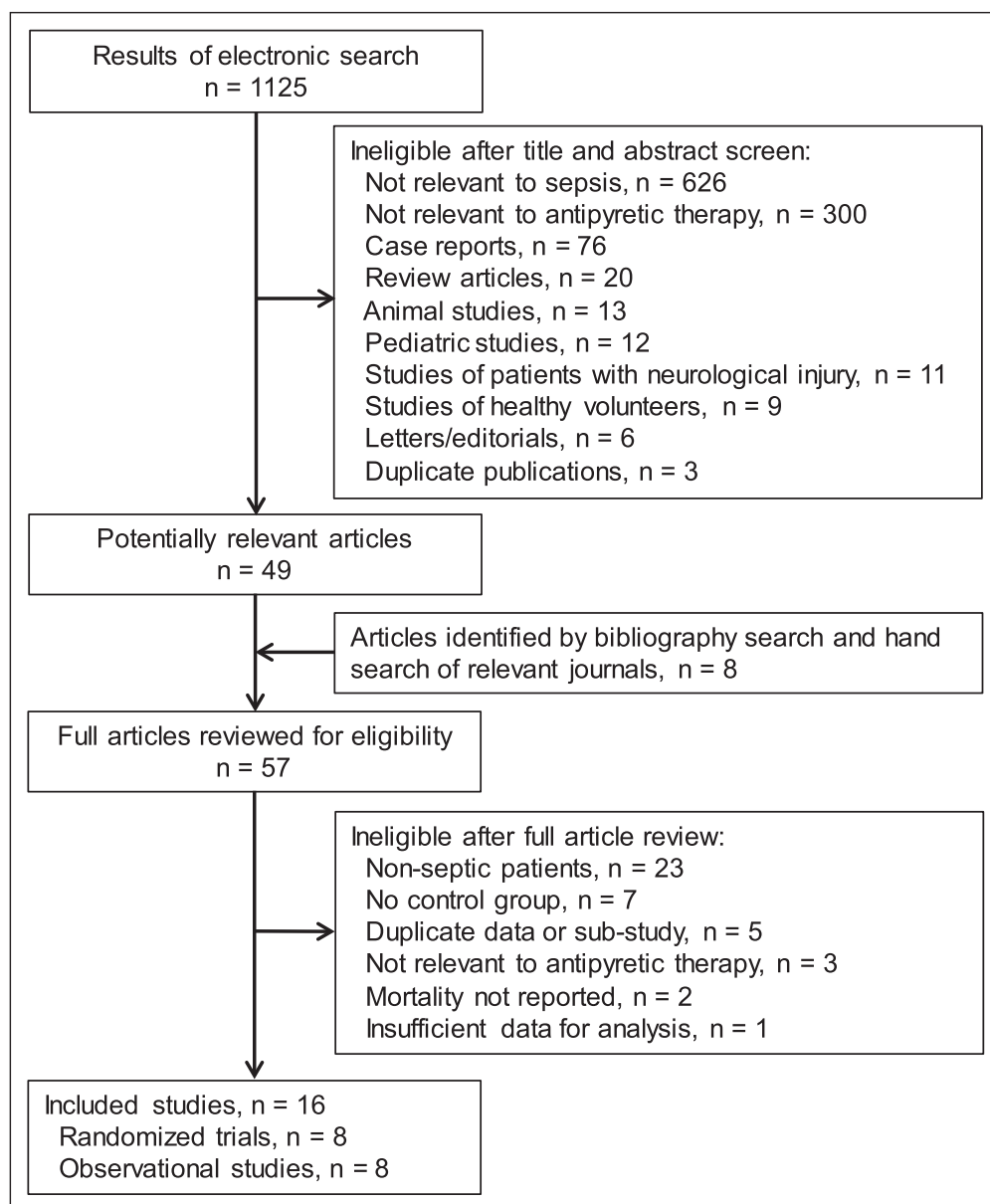


Figure 1. Flowchart of study selection.

RESULTS

Details regarding the literature search and study selection are shown in **Figure 1**. A total of 16 studies (eight randomized studies and eight observational studies) met eligibility criteria (15, 16, 26, 27, 38–49). Study characteristics are shown in **Supplemental Table 2** (Supplemental Digital Content 5, <http://links.lww.com/CCM/C425>).

Randomized Trials

The randomized studies enrolled a total of 1,531 patients (1,507 patients included in analysis of the primary outcome). Patient characteristics and outcome data for the individual trials are shown in **Supplemental Tables 3 and 4** (Supplemental Digital Content 5, <http://links.lww.com/CCM/C425>). Risk of bias assessments are shown in **Supplemental Table 5** (Supplemental Digital Content 5, <http://links.lww.com/CCM/C425>). Five studies had a low risk of bias.

Results of the meta-analyses for the primary and secondary outcomes are listed in **Table 1**. Four studies (1,198 patients) reported 28-day mortality with a pooled RR of 0.93 (95% CI, 0.77–1.13; $I^2 = 0.0\%$) comparing antipyretic therapy to control. The remaining four studies reported hospital mortality; adding this data to the analysis (1,507 total patients) resulted in a pooled RR of 0.93 (95% CI, 0.79–1.09; $I^2 = 0.0\%$) (**Fig. 2**). Subgroup analyses of 28-day/hospital mortality in febrile patients (RR, 0.96; 95% CI, 0.80–1.14; $I^2 = 0.0\%$) and patients with shock (RR, 0.91; 95% CI, 0.74–1.11; $I^2 = 0.0$) yielded similar results. Stratified analyses by type of therapy and treatment goal also did not differ significantly from that of the aggregate data (Table 1).

Analyses of secondary outcomes (Table 1) showed a significant decrease in early mortality (RR, 0.68; 95% CI, 0.49–0.92; $I^2 = 0.0\%$) with antipyretic therapy. Postintervention body temperature was also significantly lower (mean difference, -0.38°C ; 95% CI, -0.63 to -0.13 ; $I^2 = 84.0\%$) in patients treated with antipyretics. Stratified analysis of postintervention body temperature by type of intervention showed that physical cooling and nonsteroidal anti-inflammatory drugs (NSAIDs) lowered body temperature (mean difference, -0.80°C ; 95% CI, -1.06 to -0.54 and mean difference, -0.59°C ; 95% CI, -1.16 to -0.03 ; $I^2 = 85.4\%$) more effectively than acetaminophen (mean difference, -0.14°C ; 95% CI, -0.37 to 0.10 ; $I^2 = 71.3\%$). However, only one study used physical cooling as the primary antipyretic intervention. Postintervention heart rate and minute ventilation were not significantly different between the groups. The frequency of nosocomial infections and shock reversal were also unchanged (forest plots shown in **Supplemental Figs. 1–3**, Supplemental Digital Content 5, <http://links.lww.com/CCM/C425>).

Publication bias was not evident (Egger test, $p = 0.60$) (**Supplemental Fig. 4**, Supplemental Digital Content 5, <http://links.lww.com/CCM/C425>). The extended funnel plot, which graphically demonstrates the combinations of effect size and SE that would be required for an additional study to change the results of this meta-analysis to support a 28-day/hospital mortality benefit with antipyretic therapy, is shown in **Figure 3**. **Supplemental Figure 5** (Supplemental Digital Content 5, <http://links.lww.com/CCM/C425>) shows the power curve generated

from simulation-based sample size calculations. To achieve a power of 80% to change the results of this meta-analysis, an additional study would require a total sample size of approximately 29,000 patients.

Observational Studies

Eight observational studies were deemed eligible. **Supplemental Tables 6 and 7** (Supplemental Digital Content 5, <http://links.lww.com/CCM/C425>) describe the patient characteristics and results of the quality assessments. Six studies were high quality; two, low quality.

A total of 2,058 septic patients (six studies) were included in the analysis of 28-day/hospital mortality; 15,374 septic patients (two studies) were included in the analysis of early mortality. Outcome data for the individual studies, including adjusted and unadjusted ORs for mortality, are shown in **Supplemental Table 8** (Supplemental Digital Content 5, <http://links.lww.com/CCM/C425>). The pooled OR for 28-day/hospital mortality was 0.90 (95% CI, 0.54–1.51; $I^2 = 76.1\%$) (Fig. 2). The pooled OR for early mortality was 0.22 (95% CI, 0.004–13.14; $I^2 = 86.7\%$) (Supplemental Fig. 1, Supplemental Digital Content 5, <http://links.lww.com/CCM/C425>). Other secondary outcomes were not reported in a sufficient number of studies to be analyzed. No specific antipyretic method was significantly associated with mortality benefit, and stratification by study quality did not yield results that differed from the overall pooled OR (Table 1). Publication bias was not evident (Egger test, $p = 0.54$). (Supplemental Fig. 4, Supplemental Digital Content 5, <http://links.lww.com/CCM/C425>).

DISCUSSION

Despite lack of evidence showing benefit of antipyretic therapy in septic patients, treatment of fever is ubiquitous in the ICU (14). This meta-analysis was undertaken to inform clinical practice by assessing outcomes associated with antipyretic therapy. The results demonstrate that, while associated with a reduction in body temperature, antipyretic therapy does not confer a 28-day/hospital mortality benefit in septic patients. Secondary outcomes, including shock reversal and acquisition of nosocomial infections, were also unaffected by antipyretic treatments. Consistency in results was demonstrated across study design as well as in a priori subgroup and stratified analyses. Furthermore, the extended funnel plot analysis suggests that these results are likely to be robust to the impact of an additional trial in the future; none of the existing studies generated an effect size-SE combination that could change the pooled RR to favor antipyretic treatment. Additionally, simulation analysis showed that to achieve a reasonable power to change the meta-analysis results, an additional trial would need to enroll tens of thousands of patients. Based on the sample sizes and enrollment durations of the existing multicenter studies, a trial of this size seems unfeasible.

Interestingly, early mortality (occurring within 14 d or during ICU stay) was significantly lower in patients treated with antipyretic therapy in the randomized studies. This outcome was analyzed separately from 28-day/hospital mortality

TABLE 1. Summary of Meta-Analysis Results

| Analysis | Randomized Studies | | | | | Observational Studies | | | | |
|---|--------------------|-------------|--|--------|--------------------|-----------------------|-------------|-------------------|--------|--------------------|
| | No. of Studies | Sample Size | Relative Risk (95% CI) or Mean Difference (95% CI) | P | I ² , % | No. of Studies | Sample Size | OR (95% CI) | P | I ² , % |
| Primary outcome | | | | | | | | | | |
| 28-d/hospital mortality | 8 | 1,507 | 0.93 (0.79–1.09) | 0.36 | 0.0 | 6 | 2,058 | 0.90 (0.54–1.51) | 0.70 | 76.1 |
| 28-d mortality | 4 | 1,198 | 0.93 (0.77–1.13) | 0.49 | 0.0 | – | – | – | – | – |
| Febrile patients | 5 | 1,341 | 0.96 (0.80–1.14) | 0.60 | 0.0 | 4 | 1,160 | 0.59 (0.32–1.07) | 0.08 | 48.5 |
| Patients with shock | 2 ^a | 493 | 0.91 (0.74–1.11) | 0.34 | 0.0 | – | – | – | – | – |
| High-quality studies ^b | 5 | 1,238 | 0.93 (0.76–1.12) | 0.43 | 0.0 | 5 | 2,028 | 0.99 (0.62–1.60) | 0.98 | 74.0 |
| Secondary outcomes | | | | | | | | | | |
| Early mortality (≤ 14 d/ICU) | 4 | 960 | 0.68 (0.49–0.92) | 0.01 | 0.0 | 2 | 15,374 | 1.35 (1.15–1.59) | <0.001 | 0.0 |
| Frequency of nosocomial infections | 3 ^c | 684 | 1.13 (0.61–2.09) | 0.69 | 61.0 | – | – | – | – | – |
| Shock reversal | 3 | 232 | 1.13 (0.68–1.90) | 0.63 | 51.6 | – | – | – | – | – |
| Postintervention temperature (°C) | 8 | 1,510 | –0.38 (–0.63 to –0.13) | <0.003 | 84.0 | – | – | – | – | – |
| Postintervention heart rate (beats/min) | 5 | 594 | –4.2 (–10.2 to 1.9) | 0.18 | 42.9 | – | – | – | – | – |
| Postintervention minute ventilation (L/min) | 3 | 514 | –0.10 (–1.15 to 0.94) | 0.85 | 0.0 | – | – | – | – | – |
| Stratified analyses | | | | | | | | | | |
| 28-d/hospital mortality by intervention type | | | | | | | | | | |
| Acetaminophen | 3 | 753 | 0.93 (0.68–1.40) | 0.90 | 0.0 | 5 ^d | 2,028 | 0.81 (0.46–1.42) | 0.46 | 74.6 |
| NSAID | 4 | 554 | 0.94 (0.68–1.31) | 0.72 | 17.4 | 1 | 606 | 2.61 (1.11–6.12) | 0.03 | – |
| Physical cooling | 1 | 200 | 0.88 (0.65–1.19) | 0.40 | – | 2 | 636 | 0.20 (0.00–10.91) | 0.43 | 85.5 |
| 28-d/hospital mortality by treatment goal | | | | | | | | | | |
| Anti-inflammatory treatment | 5 | 594 | 0.92 (0.66–1.28) | 0.61 | 15.4 | 1 | 292 | 0.48 (0.25–0.92) | 0.03 | – |
| Fever treatment | 3 | 913 | 0.93 (0.74–1.18) | 0.55 | 0.0 | 5 | 1,766 | 1.00 (0.57–1.75) | 0.99 | 75.7 |
| 28-d/hospital mortality by treatment duration | | | | | | | | | | |
| ≥ ICU length of stay | 2 | 713 | 1.01 (0.70–1.46) | 0.95 | 0.0 | 2 | 1,287 | 1.37 (0.75–2.50) | 0.31 | 77.1 |
| < ICU length of stay | 6 | 794 | 0.91 (0.77–1.08) | 0.29 | 0.0 | 4 | 771 | 0.51 (0.26–0.98) | 0.05 | 47.6 |
| Postintervention temperature by intervention type | | | | | | | | | | |
| Acetaminophen | 3 | 756 | –0.14 (–0.37 to 0.10) | 0.26 | 71.3 | – | – | – | – | – |
| NSAID | 4 | 554 | –0.59 (–1.16 to –0.03) | 0.04 | 85.4 | – | – | – | – | – |
| Physical cooling | 1 | 200 | –0.80 (–1.06 to –0.54) | <0.001 | – | – | – | – | – | – |

NSAID = nonsteroidal anti-inflammatory drug, OR = odds ratio.

^aMemiş et al (41) excluded from analysis (100% mortality in both arms).

^bRandomized studies with a high or unclear risk of bias in less than two domains on the Cochran Collaboration Risk of Bias Tool or observational studies with a Newcastle-Ottawa Scale score greater than 7 were considered to be high quality.

^cNiven et al (42) excluded from analysis (0 nosocomial infections in septic patients in both arms).

^dThe two studies by Mohr et al (46, 47) were analyzed in the acetaminophen subgroup due to the reported low use of nonsteroidal anti-inflammatory drugs in those studies.

Dashes indicate insufficient data to analyze.

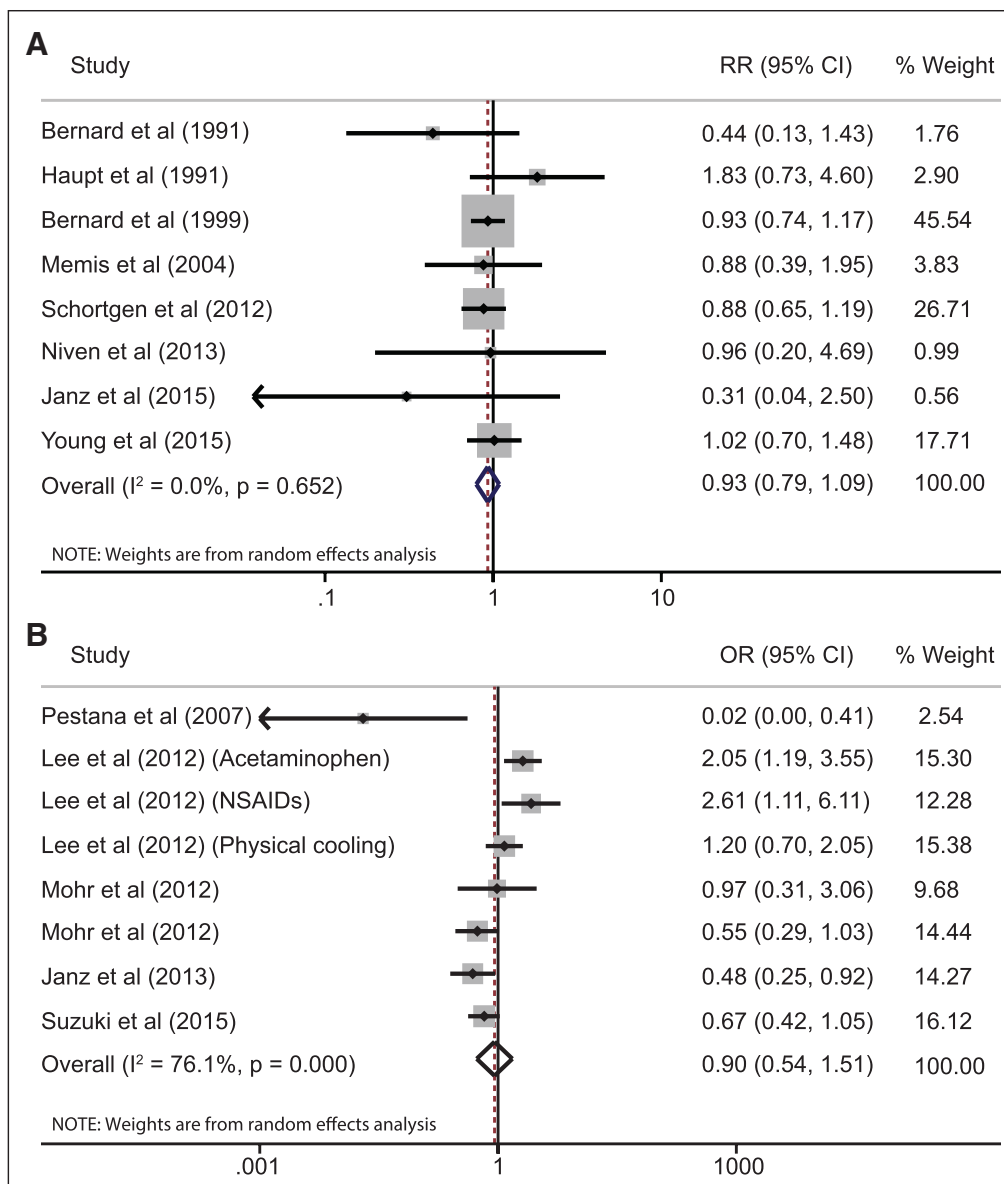


Figure 2. Results of meta-analysis for 28 days per hospital mortality in (A) randomized studies and (B) observational studies. A relative risk (RR) or odds ratio (OR) less than 1 favors antipyretic therapy. The size of the grey box corresponds to weight in the random effects analysis. NSAID = nonsteroidal anti-inflammatory drug.

because several studies reported 14-day/ICU mortality rates that differed from those at later time points (26, 27). The importance of improved early mortality, though, is questionable as a patient-centered outcome, and this finding should not influence clinical practice. One hypothesis for the decrease in early, but not later, deaths is that fever treatment blunts the immunologic benefit of hyperthermia leading to increased nosocomial infections later in the hospital course. The results of this meta-analysis demonstrated no significant differences in the acquisition of nosocomial infections among patients who did and did not receive antipyretic therapy. Analysis of this outcome, however, included only three studies, so evidence is limited. This may be an area for future study.

Proponents of fever treatment advocate that the chief benefit of antipyretic therapy in critically ill patients is a reduction

of the metabolic burden typically associated with elevated body temperature (50). This meta-analysis shows that although antipyretic therapy is effective in decreasing body temperature, heart rate and minute ventilation are less affected. Also, antipyresis did not improve mortality in the subgroup of patients with septic shock, who presumably would be the most likely to benefit from a reduction in metabolic burden. These results suggest that the potential physiologic benefit of antipyretic therapies may be overstated and does not translate into improvement in outcomes.

Definitions of fever ranged from a body temperature of 38.0°C to 38.4°C in the randomized studies and from 37.3°C to 39.5°C in the observational studies. The larger range of fever definitions in the observational studies may have contributed (along with other factors such as variations in study design, patient population, and analysis techniques) to the greater heterogeneity observed in the meta-analysis results. Of note, the observational study with the highest threshold for fever treatment (39.5°C) demonstrated the most substantial improve-

ment in 28-day/hospital mortality with antipyretic therapy (44). The implication of this finding is unclear, though, because of this study's small sample size, methodologic limitations, and unique method of physical cooling (continuous venovenous hemofiltration).

This meta-analysis has important limitations. Many of studies included in the analysis were not designed primarily to evaluate the clinical effect of fever treatment but rather the effect of specific anti-inflammatory actions of the interventions being studied. Thus, both febrile and afebrile patients were enrolled, and administration of other antipyretics beyond the specific therapy being studied was not controlled. To address this limitation, analysis of the primary outcome was stratified by the goal of the study (evaluation of anti-inflammatory treatment vs evaluation of fever treatment) and a subgroup analysis

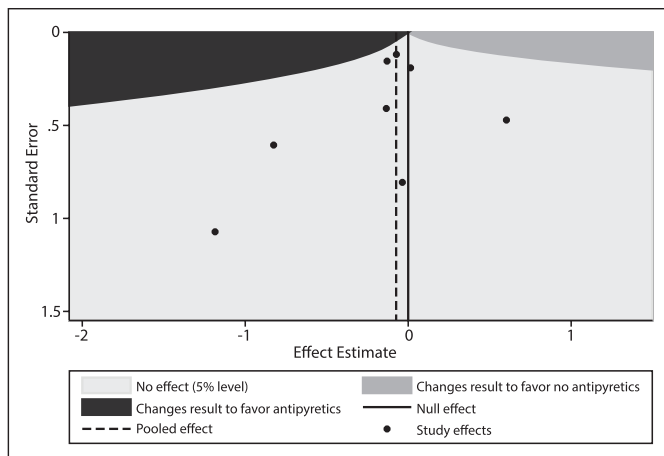


Figure 3. Extended funnel plot demonstrating the effect size and SE combinations required of an additional randomized study to change the results of this meta-analysis using a fixed effects model with an alpha level of 0.05. **Black dots** represent the effect size and SEs of the included randomized studies. To change the results of the meta-analysis to favor antipyretics, an additional study would need to have an effect estimate and SE combination that falls in the **dark grey** shaded area. Note that none of the current studies have an effect size-SE in the **dark grey** region; this suggests that the results of the meta-analysis would be robust to an additional study.

including only febrile patients was performed. The results of these analyses were virtually identical to the pooled RR and 95% CI reported for the entire cohort indicating that this limitation should not significantly affect the overall conclusions of this meta-analysis.

The studies included in this meta-analysis also varied considerably in terms of the specific antipyretic interventions being evaluated (NSAIDs, acetaminophen, physical cooling) and the duration of antipyretic treatment. Each intervention is associated with distinct effects beyond fever control which could potentially affect patient outcomes. Despite this, study heterogeneity was low ($I^2 = 0.0\%$) in the randomized studies, and stratified analysis did not show significant differences in mortality among different intervention types or intervention durations. However, since fewer studies were analyzed in each stratum as compared to the overall meta-analysis, it is possible that future studies of specific treatments would have more power to change these stratified results at any given sample size than indicated by our sample size simulation analysis.

Additionally, although the search was the most rigorous and comprehensive one on this topic to date, several studies that included mixed populations of infected and noninfected patients were excluded because mortality data could not be obtained for the subset of infected patients. Based on the extended funnel plot and the power curve generated from simulated sample size calculations, the inclusion of these studies would not have changed the results of the primary outcome.

Finally, sepsis is a heterogeneous syndrome by definition. Although our results strongly suggest that antipyretic therapy, across a broad cohort of septic patients, does not improve outcome, it is unable to inform on the individual septic patient who may accrue benefit from fever control.

CONCLUSIONS

Antipyretic treatment does not significantly improve 28-day/hospital mortality in adult patients with sepsis. Additional studies are unlikely to be powered sufficiently to change this conclusion.

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