### **Original Article**

# Early growth characteristics and the risk of reduced lung function and asthma: a meta-analysis of 25,000 children

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ABSTRACT 1 2 Background Children born preterm or with a small-size-for-gestational-age are at increased 3 risk for childhood asthma. We tested the hypothesis that these associations are explained by 4 reduced airway patency. Methods and Materials We used individual participant data of 24,938 children from 24 birth 5 6 cohorts to examine the associations of gestational age, size-for-gestational-age, and infant 7 weight gain with childhood lung function (forced expiratory volume in 1 second (FEV<sub>1</sub>), 8 FEV<sub>1</sub>/forced vital capacity (FEV<sub>1</sub>/FVC), forced expiratory flow after exhaling 75% of the vital 9 capacity (FEF<sub>75</sub>)) and risk of doctor-diagnosed asthma (age range 3.9 – 19.1 years). 10 Second, we used similar models to explore whether these lung function outcomes mediated 11 the associations of early growth characteristics with the risk of childhood asthma. 12 Results Children born with a younger gestational age had a lower FEV<sub>1</sub>, FEV<sub>1</sub>/FVC, and FEF<sub>75</sub>, whereas those born with a smaller size-for-gestational-age at birth had lower FEV<sub>1</sub> 13 14 but higher FEV<sub>1</sub>/FVC (p-values<0.05). Greater infant weight gain was associated with higher FEV<sub>1</sub>, but lower FEV<sub>1</sub>/FVC and FEF<sub>75</sub> in childhood (p-values<0.05). All associations were 15 16 present across the full ranges and independent of other early life growth characteristics. In 17 line with these observations, preterm birth, low birth weight and greater infant weight gain were associated with an increased risk of childhood asthma (pooled odds ratio (95% CI): 18 19 1.50 (1.2, 1.57), 1.40 (1.18, 1.65) and 1.23 (1.17, 1.28), respectively). These associations 20 were mediated up to 45% by lung function characteristics. Mediation analyses suggested that 21 FEV<sub>1</sub>, FEV<sub>1</sub>/FVC and FEF<sub>75</sub> may explain 6.3 (2.5, 9.9)% to 44.6 (14.6, 81.1)% of the 22 associations between early growth characteristics and lung function. 23 **Conclusions** Results from this individual participant meta-analysis suggest that younger 24 gestational age at birth, smaller size-for-gestational-age at birth, and greater infant weight 25 gain across the full ranges were associated with adaptations in childhood lung function, and 26 that these associations explain to a substantial extent the risk of childhood asthma.

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28

# 29 INTRODUCTION

30	Children born extremely preterm or with a low birth weight have high rates of neonatal
31	respiratory diseases such as infant respiratory distress syndrome and bronchopulmonary
32	dysplasia (1). An accumulating body of evidence suggests that these children also have an
33	increased risk of chronic obstructive respiratory diseases in adulthood (2). More recent,
34	prospective studies in children suggest that preterm birth and small size for gestational age
35	at birth increase the risk of childhood asthma (3). Recent results of a meta-analysis of
36	individual participant data of 147,000 children participating in prospective birth cohort studies
37	showed consistent associations of younger gestational age at birth and greater infant weight
38	gain with childhood asthma (4). The associations of lower birth weight with childhood asthma
39	seem to be largely explained by gestational age at birth (4). The mechanisms underlying the
40	associations of early growth characteristics with childhood asthma are not known yet. Airway
41	caliber is a key determinant of total airway resistance. A reduced airway caliber could result
42	in airway obstruction that predisposes to asthma and chronic obstructive pulmonary
43	diseases (5-7). Therefore, we hypothesized that the associations of early growth
44	characteristics with childhood asthma might be explained by , but might include
45	developmental adaptations of the lungs and airways, leading to relatively small airways and,
46	hence, a reduction in expiratory flows reflected by lower lung function values (8). Thus far,
47	previous studies focused on the associations of birth weight and infant weight gain with
48	childhood lung function have reported inconsistent results (9-16). These inconsistent results
49	might be due to the different ages at which spirometry was performed, and not taking other
50	early growth characteristics or potential confounders into account.
51	To test the hypothesis that the associations of early life growth characteristics with
52	childhood asthma are explained by reduced airway patency, we performed an individual
53	participant data meta-analysis of 24,938 children from 24 birth cohort studies. We examined
54	the strength, consistency, and independence of the associations of gestational age at birth,
55	birth weight and infant weight gain with lung function outcomes in childhood and whether

these lung function outcomes explain the previously reported associations of early growth

- 57 characteristics with risk of childhood asthma.
- 58
- 59 METHODS
- 60

61 Sources of data

62 European population-based birth- and mother-child cohorts participated if they included 63 children born between 1989 and 2011, had information available on at least gestational age 64 and weight at birth and lung function measurements in childhood (until age 18 years), and 65 were willing and able to exchange original data.(4) We identified 52 European cohorts 66 selected from existing collaborations on childhood health or asthma-related outcomes 67 (www.chicosproject.eu, www.birthcohortsenrieco.net, www.ga2len.org, 68 and www.birthcohorts.net; assessed until May 29, 2012). In total, 24 cohorts, comprising 69 data on 24,938 children, fulfilled the criteria (Supporting Information: S-figure 1). 70 Information about gestational age and weight at birth and weight in the first year of 71 life was obtained by measurements, medical registries or parental questionnaires 72 (Supporting Information: S-table 1). We created gestational age-adjusted birth weight 73 standard deviation scores (birth weight SDS) based on European reference values (17). 74 Infant weight gain in the first year was defined as the difference between weight at age 1 75 year (range 6-18 months) and weight at birth, divided by the number of months between 76 these two measurements. Standard deviation scores (SDS) for age-specific infant weight 77 gain were derived by intra-cohort means and standard deviations (18). Cohort specific 78 growth characteristics are given in the Supporting Information (S-table 2). 79 All cohorts obtained lung function measurements by spirometry, most of them (n =80 21) of which 22 according to the recent guidelines of the American Thoracic Society / 81 European Respiratory Society (ATS/ERS) (19-21), and 2 according to earlier guidelines of 82 the ATS (22) or ERS and European Coal and Steel Community (23) (Supporting 83 Information: S-table 1) (21-23). If cohorts had collected lung function data at multiple time

**Field Code Changed** 

84	points (n = 6 cohorts), we used the measurement closest to the mean age of children (8.5
85	years) in the full meta-analysis. Variables for analyses were forced vital capacity (FVC),
86	forced expiratory volume in 1 second (FEV1), forced mid-expiratory flow (FEF25-75) and
87	forced expiratory flow after exhaling 75% of the vital capacity (FEF $_{75}$ ). We mainly focused on
88	$FEV_1$ , $FEV_1$ /FVC, and $FEF_{75}$ , which reflect reduced airway patency in obstructive lung
89	diseases such as asthma or bronchopulmonary dysplasia due to preterm birth or low birth
90	weight (24, 25). All lung function variables were converted into sex-, height-, age-, and
91	ethnicity (Caucasian versus non-Caucasian) -adjusted Z-scores based on the Global Lung
92	Initiative reference values (26). Asthma (yes / no) was defined as ever physician diagnosed
93	asthma, and was obtained by medical registries (2 cohorts) or parental questionnaires
94	adapted from the International Study on Asthma and Allergy in Childhood (ISAAC) (27) (22
95	cohorts) at the age of spirometry (S-table1). Cohort specific characteristics of lung function
96	measurements and asthma are given in the Supporting Information (S-table 3).
97	We included covariates based on known associations with childhood lung function
98	from previous studies (28, 29). Information on covariates was mainly assessed by
99	questionnaires (Supporting Information: S-table 1). Potential confounders included
100	maternal educational level, smoking during pregnancy, smoking during infancy of their
101	offspring, history of asthma or atopy, child's sex, siblings, day care attendance in the first 2
102	years of life, breastfeeding, lower respiratory tract infections in the first 2 years of life,
103	eczema, inhalant allergies, and body mass index (BMI) at the moment of lung function
104	measurement. Cohort specific characteristics of all covariates are given in the Supporting
105	Information (S-tables 4-5).
106	
107	Statistical analysis
108	First, we conducted 1-stage random effect regression analyses to study the separate and
109	combined associations of gestational age, birth weight and infant weight gain with $FEV_1$ ,
110	FVC, FEV <sub>1</sub> /FVC, FEF <sub>25-75</sub> and FEF <sub>75</sub> . For these analyses, individual participant data from all
111	cohorts were combined and modeled simultaneously taking into account clustering of

111 cohorts were combined and modeled simultaneously taking into account clustering of

112 participants within studies (30). To prevent multicollinearity in our regression models, we 113 started to assess the associations of gestational age and birth weight with lung function 114 separately. Thereafter, we assessed whether the associations of birth weight with lung 115 function was driven by gestational age by creating gestational age adjusted birth weight 116 standard deviation scores. The models focused on the associations of infant weight gain 117 with lung function outcomes were adjusted for gestational age and weight at birth. To test 118 non-linear and dose-response associations, we categorized gestational age, birth weight 119 SDS and infant weight gain SDS. As a sensitivity analysis, we conducted a 2-stage random 120 effect meta-analysis to study the associations of gestational age, birth weight, and infant 121 weight gain, and dichotomized preterm birth and low birth weight with each lung function 122 outcome. For this analysis, we used linear regression models per cohort, after which pooled 123 regression coefficients ( $\beta$ 's) from the per cohort effect estimates were calculated. We tested for heterogeneity between effect estimates using l<sup>2</sup> (31, 32). For all analyses, the first model 124 125 was adjusted for child's sex (crude model), the second model was additionally adjusted for 126 potential confounders (full model). To determine interactive effects between gestational age, 127 birth weight and infant weight gain we added these terms multiplicative in the full model. 128 Since we used Northern-European reference curves for birth weight SDS, we performed a 129 sensitivity analysis to explore whether the associations were different in North-Western 130 European subjects only. Numbers were too small to perform these analyses separately in 131 other European regions. To assess differences in results related to pubertal growth changes, 132 we repeated our analyses is strata of children aged < 11 years and ≥11 years (33). We also 133 performed a complete-case sensitivity analysis to explore any differences between complete 134 and non-complete-case analyses, and sensitivity analyses in which we excluded cohorts that 135 used parental report of early growth characteristics or that did not perform spirometry 136 measurements according to the ATS/ERS guidelines. 137 Second, we conducted a 1-stage random effect regression analysis to assess the

associations of early growth characteristics with asthma, and observed whether changes in
the effect estimates occurred after additional adjustment for lung function measures (FEV<sub>1</sub>,

140	FVC, FEV <sub>1</sub> /FVC, FEF <sub>25-75</sub> and FEF <sub>75</sub> ) as potential mediators (mediator model). The
141	difference between the original effect estimates and the effect estimates after additional
142	adjustment for potential mediators was expressed as percentage change. The percentage
143	change of the effect estimate was calculated by the formula: 100 x (effect estimate $_{mediator}$ -
144	effect estimate <sub>model 1 original model</sub> )/( effect estimate <sub>model 1 original model</sub> - 1). A 95% confidence interval
145	for the percentage change of the effect estimate was calculated using a bootstrap method
146	with 1,000 resamplings (34-36).
147	For all analyses, missing values in covariates were used as an additional group in
148	the categorical variables to prevent exclusion of non-complete cases. Statistical analyses
149	were performed with R version 3.0.0 (libraries rmeta and metafor; The R foundation for
150	Statistical Computing), and Comprehensive Meta-Analysis (Biostat, US).
151	
152	RESULTS
153	
154	Subject characteristics
155	Information about the main characteristics of the cohorts are given in Table 1. Detailed
156	information about determinants, outcomes and covariates is given in the Supporting
157	Information (S-tables 1-5). Of all participants, 8.2% (n = 2,053) was born preterm (<37
158	weeks of gestational age), and 4.8% (n = 1,191) was born with a low birth weight (<2,500 $$
159	gram). The mean age at which spirometry assessments were performed was 8.5 (range 3.9
160	- 19.1) years. <u>The proportion of children aged ≥11 years was 11.9% (n = 2,972).</u>
161	
162	Early growth measures and lung function outcomes
163	Results from the 1-stage random effect models showed that younger gestational age at birth
164	was, across the full range, associated with lower FEV1, FEV1/FVC and FEF75 in childhood
165	
	(p-values for trend <0.01) (Figures 1A-C). A smaller size-for-gestational-age at birth across
166	(p-values for trend <0.01) ( <b>Figures 1A-C</b> ). A smaller size-for-gestational-age at birth across the full range was associated with lower $FEV_1$ and higher $FEV_1/FVC$ (p-values for trend

168	(Figure 1F). Greater infant weight gain was associated with a higher $FEV_1$ , but with a lower
169	FEV <sub>1</sub> /FVC and FEF <sub>75</sub> (p-values for trend <0.01; Figures 1G-I). Most associations showed a
170	linear trend, except for the associations of birth weight with FEV <sub>1</sub> /FVC and infant weight gain
171	with FEV <sub>1</sub> and FEV <sub>1</sub> /FVC which were non-linear (Figures 1E, G, H).
172	To explore the combined effects of gestational age, birth weight SDS and infant
173	weight gain SDS, we performed tests for interaction between these early growth
174	characteristics. These tests for interaction were-only significant for gestational age and birth
175	weight SDS in relation to <del>lung function outcomes<u>FEV1</u>, FEV<u>1</u>/FVC, FEF<sub>25-75</sub> and FEF<sub>75</sub> (p-</del>
176	values for interaction <0.01; <b>Figure 2<u>, S-table 9</u>)</b> . <del>Specifically,<u>Stratified analyses showed</u></del>
177	that a lower birth weight was associated with lower $FEV_1$ and $FEV_1/FVC$ among children
178	born after $\geq$ 32 weeks only, whereas higher birth weight was associated with FEF <sub>75</sub> only
179	among term born children (p-values for strata <0.05).
180	No differences in results were observed when we used 2-stage random effect
181	models of combined effect estimates (Supporting Information: S-tables 6-7). Also, the
182	results from the sensitivity analyses showed similar results when we used cohorts with
183	North-Western European subjects only, when we excluded cohorts that did not perform
184	spirometry measurements according to the recent ATS/ERS guidelines, when we performed
185	stratified analyses for children aged < 11 years or ≥ 11 years ( <b>S-table 8</b> ), or when we
186	excluded cohorts that used parental report of early growth characteristics (data not shown).
187	Also, the results were similar when we used cohorts with North-Western European subjects
188	only, complete-cases, or when we excluded cohorts that used parental report of early growth
189	characteristics (Supporting Information: S-table 8) or did not perform spirometry
190	measurements according to the ATS/ERS guidelines (data not shown).
191	Figure 3 shows that compared to term born children, those born preterm had a lower
192	$FEV_1$ , $FEV_1$ /FVC and $FEF_{75}$ , (pooled Z-score (95% CI): -0.20 (-0.26, -0.14), -0.15 (-0.21, -0.21))
193	0.09) and -0.19 (-0.27, -0.11), respectively). Also, compared to normal birth weight children,
194	those with a low birth weight had lower FEV <sub>1</sub> , FEV <sub>1</sub> /FVC and FEF <sub>75</sub> (-0.29 (-0.38, -0.21) and
195	-0.16 (-0.25, -0.08) and -0.17 (-0.26, -0.08) respectively), independent of gestational age.

196	Results of associations of growth characteristics with all lung function outcomes, including
197	FVC and FEF <sub>25-75</sub> are given in the <b>Supporting Information: S-tables 6-8</b> .
198	
199	Early growth, lung function and asthma
200	Preterm birth, low birth weight and greater weight gain were all associated with an increased
201	risk of childhood asthma (OR (95% CI): 1.50 (1.2, 1.57), 1.40 (1.18, 1.65) and 1.23 (1.17,
202	1.28), respectively. The associations of preterm birth and low birth weight attenuated after
203	additional adjustment for FEV <sub>1</sub> , FEV <sub>1</sub> /FVC or FEF <sub>76</sub> . Mediation analyses suggested that
204	FEV <sub>1</sub> , FEV <sub>1</sub> /FVC and FEF <sub>75</sub> may explain 6.3 (2.5, 9.9)% to 44.6 (14.6, 81.1)%. Specifically,
205	after additional adjustment for FEV1, FEV1/FVC or FEF75, the associations of preterm birth
206	with asthma attenuated with -7.3 (-18.8, -0.9)%, -14.4 (-39.6, -2.8)% and -39.0 (-69.3, -
207	3.4)%, respectively. Similarly, the associations of low birth weight with asthma attenuated
208	with -19.0 (-37.3, -11.8)%, -21.6 (-47.3, -11.4)% and -21.6 (-47.3, -11.4)%, respectively
209	( <b>Table 2</b> ). The strongest mediating effect was observed for $\text{FEF}_{75}$ for the association
210	between gestational age and asthma (44.6 (81.1, 14.6)%). Overall, the strongest
211	attenuations towards non-significant were observed after adjustment for FEF75 with an up to
212	45% reduction of the effect estimate (Table 2). A similar trendSimilar trends were was
213	observed for greater weight gain, although the associations did not attenuate into non-
214	significant.
215	
216	DISCUSSION
217	In this meta-analysis of individual participant data of 24,938 children from 24 birth cohorts,
218	we observed that lower gestational age, smaller size at birth and greater infant weight gain
219	were all associated with lower childhood FEV1. The positive associations of birth weight and
220	infant weight gain with FVC were larger than of the positive associations of birth weight and
221	infant weight gain with FEV <sub>1</sub> . This combination resulted in associations of higher birth weight
222	and infant weight gain with lower FEV1/FVC. Also, a lower gestational age at birth was
223	associated with a lower $FEF_{75}$ in childhood, suggesting persistent reduction of small airways

224	patency. A greater infant weight gain was associated with lower FEF75Remarkably, these
225	associations were present across the full-range of early growth and not restricted to clinically
226	diagnosed preterm- or low birth weight children. Also, the observed associations of the early
227	life growth characteristics with lung function outcomes were independent of each other.
228	Stratified analyses showed that children born very preterm with a relatively low birth weight
229	had the lowest $FEV_1$ and $FEV_1/FVC$ . The associations of early growth characteristics with
230	childhood asthma were partly explained by lung function adaptations.

231 Whereas lung growth continues until the early adulthood, the most rapid 232 development of airways and alveoli occurs in early life (37). Developmental adaptations in 233 fetal life and infancy due to early life adverse exposures might result in impaired lung growth 234 with smaller airways, decreased lung volume, and subsequently to an increased risk of 235 bronchopulmonary dysplasia, asthma or COPD (9, 14, 38). Previous studies suggest that 236 children with asthma already have a reduced lung function in the first months of life, and that 237 this deficit progresses into childhood and early adulthood (39, 40). Airway caliber is a key 238 determinant of total airway resistance and reduced caliber is a prominent feature of asthma 239 and chronic obstructive pulmonary diseases (5-7). Lower lung function in early life is likely to 240 lead to lower peak lung function in early adulthood, and the natural decline in FEV<sub>1</sub> from that 241 point onwards will be accelerated by any additional adverse exposures (41). Thus, lung 242 function during the lifecourse seems to be programmed at least partly in early life. 243 Children born preterm or with a very low birth weight are at increased risk of neonatal 244 respiratory diseases (1). We observed that children born at a younger gestational age had a 245 lower FEV<sub>1</sub>, even after taking FVC into account, and a lower FEF<sub>75</sub> in childhood. These 246 associations were not only present among children born very preterm, but across the full 247 range of gestational age at birth. Moreover, the associations of preterm birth with childhood

asthma were partly explained by lung function. These findings are in line with previous

249 studies showing persistent lung function adaptions in children and adults born preterm. A

recent meta-analysis of 28 published studies showed that children born between 24 and 36

251	weeks had a lower $FEV_1$ at ages 5 up to 23 years (42). These and other studies suggest that
252	preterm birth has adverse effects on lung function, persisting into adulthood (42-44).
253	In the present study, a lower birth weight was associated with lower $FEV_1$ in
254	childhood. This suggests that a lower birth weight leads to a persistent reduction of airway
255	patency. A previous study analyzed 10 studies examining the associations of birth weight
256	with $FEV_1$ in adults (range 19 – 70 years) (10). The authors reported a modest positive
257	association between $FEV_1$ and birth weight. Two recent studies from longitudinal birth
258	cohorts among adults reported strong positive associations of birth weight with $FEV_1$ and
259	$\text{FEF}_{25-75}$ in young adults aged 21 and 31 years (9, 11). The effect of birth weight was
260	independent of preterm birth in both studies. However, studies among children showed
261	conflicting results (12, 13). We observed an association of lower birth weight with lower
262	$FEV_1$ , independent of gestational age at birth. We previously reported that the effect of lower
263	birth weight on asthma was largely explained by gestational age (4). Therefore, although
264	gestational age-adjusted birth weight is associated with lower lung function this seems not
265	related to the risk of clinically manifest childhood asthma.
265 266	related to the risk of clinically manifest childhood asthma. Previous studies examining associations between infant weight gain and childhood
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266 267	Previous studies examining associations between infant weight gain and childhood lung function have reported inconsistent results (14-16). Differences might be due to
266 267 268	Previous studies examining associations between infant weight gain and childhood lung function have reported inconsistent results (14-16). Differences might be due to different ages at which spirometry was performed, not taking other weight characteristics
266 267 268 269	Previous studies examining associations between infant weight gain and childhood lung function have reported inconsistent results (14-16). Differences might be due to different ages at which spirometry was performed, not taking other weight characteristics into account, such as birth weight or current body mass index, and possible hidden bias due
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266 267 268 269 270 271 272 273 273	Previous studies examining associations between infant weight gain and childhood lung function have reported inconsistent results (14-16). Differences might be due to different ages at which spirometry was performed, not taking other weight characteristics into account, such as birth weight or current body mass index, and possible hidden bias due to the use of mL instead of Z-scores for lung function (45). In line with the findings for birth weight, we observed that lower infant weight gain was associated with a lower childhood FEV <sub>1</sub> . This association was fully explained by FVC. <u>These results suggest dysanapsis, in</u> which airways remained small in relation to total lung volume as a result of a mismatch <u>between airway and alveolar growth (46). Hence, a lower FEV<sub>1</sub> due to lower birth weight or</u>
266 267 268 269 270 271 272 273 274 275	Previous studies examining associations between infant weight gain and childhood lung function have reported inconsistent results (14-16). Differences might be due to different ages at which spirometry was performed, not taking other weight characteristics into account, such as birth weight or current body mass index, and possible hidden bias due to the use of mL instead of Z-scores for lung function (45). In line with the findings for birth weight, we observed that lower infant weight gain was associated with a lower childhood FEV <sub>1</sub> . This association was fully explained by FVC. These results suggest dysanapsis, in which airways remained small in relation to total lung volume as a result of a mismatch between airway and alveolar growth (46) <sub>2</sub> Hence, a lower FEV <sub>1</sub> due to lower birth weight or infant weight can be explained by a lower lung volume. Greater infant weight gain was also

279	inflammatory effects in the airways (49), or a direct effect of increased body weight on lung
280	function (50). However, our analyses were adjusted for childhood body mass index. Further
281	studies are needed to explore whether the associations of infant weight gain with end-
282	expiratory flows are explained by specific adiposity-related measures or biomarkers.
283	To the best of our knowledge this is the first study that examines the individual and
284	combined associations of the main early growth characteristics with childhood lung function
285	outcomes, and whether lung function adaptations explain the previously reported
286	associations of early growth characteristics with childhood asthma. showing associations of
287	all early growth characteristics with lung function and subsequent risk of childhood asthma.
288	Our results suggest that respiratory consequences of preterm birth and a low birth weight
289	present across the full range. This observation might have important population effects,
290	since the largest majority of children are in the less extreme ranges of gestational age and
291	weight at birth. Furthermore, our results suggest that the associations of gestational age,
292	birth weight and infant weight gain with childhood asthma are at least partly explained by
293	adaptions in airway caliber. We observed strong effect estimates with wide confidence
294	intervals. Therefore, these mediation effects should be interpreted carefully. The effect
295	estimates for the observed associations could be considered as small and without clinical
296	relevance for individuals. However, the associations may be important from an etiological
297	respiratory developmental perspective and may be important on a population-level. The
298	associations of early growth characteristics with lung function outcomes seemed already
299	established before the pubertal growth spurt. The largest lung and airway growth occurs
300	before pubertal growth spurt (37, 51), with FVC increasing proportionately more than the
301	FEV <sub>1</sub> (33). Lung and airway growth is proportionally less after start of the pubertal growth
302	spurt (33), which might explain the similar effect estimates before and after the pubertal
303	growth spurt. Further studies are needed to identify the developmental adaptations of the
304	lungs and immune system that might explain the mediating effect of lung function on the
305	associations of early growth characteristics with childhood asthma. More information is
306	needed on the specific maternal and childhood exposures, which lead to the specific early
	14

307	growth patterns and affect the risk of later life respiratory diseases. Identification of
308	modifiable exposures may lead to development of future preventive strategies.
309	Some methodological limitations need to be discussed. We used data from 24
310	ongoing cohort studies. Missing values always occur in these studies. Since we did not have
311	additional data on patterns of missing values in all 24 cohorts, we were not able to perform
312	multiple imputation. Data on childhood asthma was mainly obtained by parental
313	questionnaires adapted from the International Study on Asthma and Allergy in Childhood
314	(ISAAC) (27). This questionnaire has been validated in various age groups in many
315	countries against measurements of bronchial hyperresponsiveness and doctor-diagnosed
316	asthma, and is widely accepted in epidemiological studies. We did not have information on
317	use of asthma medication, which might have influenced the lung function values in asthmatic
318	patients. This missing information on asthma medication may have influenced our effect
319	estimates. We would expect that asthmatic children who use asthma medication would in
320	general have had a higher lung function values in case of good adherence and inhaler
321	technique. We used GLI reference data to convert lung function values into Z-scores. These
322	prediction equations were based on 74,187 individuals including 31,840 individuals aged
323	<20 years, of whom 58% were assessed before, and 42% were assessed during pubertal
324	growth spurt (26). To date, the GLI normal values are considered the most accurate
325	reference values for all age ranges, and have been adopted by both the ATS and ERS. For
326	the covariates, we imputed missing values as additional category to prevent exclusion of
327	non-complete cases. No differences in results were observed in complete case analyses. <u>No</u>
328	direct clinical and laboratory information about pubertal growth was available. Also,
329	although we took major potential confounders into account, residual confounding may still be
330	an issue. No information was available about e.g. exposure to environmental micro-
331	organisms or asthma severity. Exploring mediation of lung function for the association of
332	early growth characteristics with asthma using the method proposed by Baron and Kenny
333	might have been limited by misclassification of lung function measurements or asthma
334	diagnosis although we aimed to reduce this issue by multi-level modelling (52). Most of the

335	participating studies had measured childhood lung function and asthma at the same age.
336	Therefore, further follow-up studies with longitudinally measured detailed data on lung
337	function and asthma or related symptoms from birth onwards are needed to disentangle the
338	direction of causality.
339	In conclusion, younger gestational age, lower birth weight and lower infant weight
340	gain were independently associated with persistent changes in childhood lung function.
341	These associations were present across the full spectrum of these early growth
342	characteristics. Stratified analyses showed that children born very preterm with a relatively
343	low birth weight had the lowest FEV <sub>1</sub> and FEV <sub>1</sub> /FVC. Our results suggest that associations
344	of early growth with the risk of childhood asthma were partly explained by lung function
345	adaptations. Thus, fetal and infant growth patterns may persistently affect lung function, and
346	thereby contribute to the risk of respiratory diseases in later life.
347	
348	Author's contributions

- 349 HD, AS, JJ, VJ, and LD contributed to the study design, data analysis plan, data collection,
- 350 data analysis, data interpretation, writing, reviewing the manuscript critically and gave
- 351 consent for submission. All other authors contributed equally to study design, data analysis
- 352 plan, data collection, reviewing the manuscript critically and gave consent for submission.

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**Figure 1.** Associations of gestational age, birth weight and infant weight gain with  $FEV_1$ ,  $FEV_1/FVC$  ratio and  $FEF_{75}$ .

## Legend:

Values represent Z-scores differences (95% confidence interval) from multi-level random effect models for the associations of gestational age at birth (A, B, C), gestational age adjusted birth weight (birth weight SDS) (D, E, F), and infant weight gain (SDS) (G, H, I) with lung function outcomes, compared with reference groups. Reference groups were 40-42.9 weeks of gestational age, 0-0.99 birth weight SDS and 0.00 – 0.99 infant weight gain (SDS) (largest groups), and represented by an open bullet. Lung function outcomes are forced expiratory volume in 1 second (FEV<sub>1</sub>), FEV<sub>1</sub>/forced vital capacity (FVC) ratio, and forced expiratory flow at 75% of the exhaled FVC (FEF<sub>75</sub>). Models are adjusted for maternal education, smoking during pregnancy, smoking during childhood, atopy, asthma and child's sex, number of siblings, daycare attendance, breastfeeding, respiratory tract infections, childhood eczema, inhalant allergies and body mass index. Infant weight gain SDS was additionally adjusted for birth weight and gestational age at birth.

**Figure 2.** Combined associations of gestational age and birth weight with FEV<sub>1</sub>, FEV<sub>1</sub>/FVC ratio and FEF<sub>75</sub>.

#### Legend:

Values are Z-score differences (95% confidence interval) from multi-level models for the combined associations of gestational age at birth and birth weight SDS (A, B, C) with lung function outcomes, compared with reference groups. Reference groups were >37 weeks of gestational age with -1.00 to 0.99 birth weight SDS (largest group), and represented by a bullet. Lung function outcomes are forced expiratory volume in 1 second (FEV<sub>1</sub>), FEV<sub>1</sub>/forced vital capacity (FVC) ratio, and forced expiratory flow at 75% of the exhaled FVC (FEF<sub>75</sub>). Models are adjusted for maternal education, smoking during pregnancy, smoking

during childhood, atopy, asthma and child's sex, number of siblings, daycare attendance, breastfeeding, respiratory tract infections, childhood eczema, inhalant allergies and body mass index. \*P-value < 0.05. \*\*P-value < 0.01. Given p-values reflect differences between birth weight SDS groups (A, B, C) within strata of gestational age using -1.00 to 0.99 birth weight SDS as reference group. P<sub>int</sub>: p-values of multiplicative interaction terms.

**Figure 3.** Forest plots of the associations between preterm birth and low birth weight with  $FEV_1$ ,  $FEV_1/FVC$  ratio and  $FEF_{75}$ .

# Legend:

Values are pooled Z-score differences (95% confidence interval) from random effect metaanalysis for the associations of preterm birth vs. term birth (A, B, C) and low birth weight vs. normal birth weight (D, E, F) with lung function outcomes. Lung function outcomes are forced expiratory volume in 1 second (FEV<sub>1</sub>), FEV<sub>1</sub>/forced vital capacity (FVC) ratio, and forced expiratory flow at 75% of the exhaled FVC (FEF<sub>75</sub>). Models are adjusted for maternal education, smoking during pregnancy, smoking during childhood, atopy, asthma and child's sex, number of siblings, daycare attendance, breastfeeding, respiratory tract infections, childhood eczema, inhalant allergies and body mass index. Low birth weight was adjusted for gestational age.

Cohort name (country)	N	Birth years	Gestational age at birth (weeks)	Birth weight (gram)	FVC	FEV <sub>1</sub>	FEV₁/ FVC	FEF <sub>25-75</sub>	FEF <sub>75</sub>	Childhood asthma
			Median (5-95% range)	Mean (SD)	Mean Z-score (SD)	Mean Z-score (SD)	Mean Z-score (SD)	Mean Z-score (SD)	Mean Z-score (SD)	Yes
ALSPAC (United Kingdom)	6,873	1991- 1992	39.5 (1.9)	3,424 (543)	0.49 (1.28)	0.44 (1.17)	-0.07 (1.15)	0.04 (1.08)	0.30 (1.06)	17.9 (1,231)
BAMSE (Sweden	2,042	1994- 1996	39.9 (1.8)	3,537 (551)	0.65 (0.93)	0.45 (0.96)	-0.37 (0.89)	-	-	14.8 (303)
BILD (Switzerland)	159	1999- ongoing	39.7 (1.3)	3,367 (441)	-0.23 (0.98)	0.02 (0.89)	0.33 (0.95)	-0.06 (0.87)	-	-
CONER (Italy)	217	2004- 2005	39.2 (1.4)	3,335 (457)	-1.76 (0.82)	-1.04 (0.90)	0.51 (1.65)	0.45 (1.00)	-	6.0 (13)
COPSAC2000 (Denmark)	314	1998- 2001	40.0 (1.6)	3,529 (531)	-0.53 (0.98)	-0.11 (1.03)	0.47 (0.95)	-	-	18.8 (59)
EDEN (France)	897	2003- 2005	39.3 (1.7)	3,284 (514)	-1.08 (1.05)	-0.77 (1.03)	0.21 (0.97)	-0.39 (1.01)	0.16 (0.88)	18.1 (162)
GASPII (Italy)	453	2003- 2004	39.2 (1.8)	3,314 (530)	0.06 (0.76)	-0.01 (0.88)	-0.15 (0.97)	-0.30 (0.90)	-	6.6 (30)
GENERATION R (The Netherlands)	1,927	2002- 2006	39.7 (1.9)	3,392 (576)	0.23 (0.92)	0.15 (0.95)	-0.19 (0.92)	0.15 (1.05)	-0.09 (0.89)	5.5 (106)
GENERATION XXI (Portugal)	1,562	2005- 2006	38.4 (2.1)	3,152 (551)	0.41 (0.95)	0.59 (0.98)	0.21 (0.82)	0.12 (0.85)	0.44 (0.80)	6.5 (102)
GINI (Germany)	707	1995- 1998	-	3,493 (479)	-	0.02 (0.92)	-	-	-	5.9 (49)
INMA Gipuzkoa (Spain)	277	2006- 2008	39.7 (1.4)	3,284 (436)	-0.54 (1.16)	-0.59 (1.17)	-0.05 (0.91)	-0.45 (0.99)	-0.16 (1.00)	5.4 (15)
INMA Menorca (Spain)	367	1997- 1998	39.2 (1.8)	3,200 (493)	0.01 (1.13)	-0.16 (1.07)	-0.24 (1.19)	-0.42 (1.29)	-0.06 (1.32)	4.9 (18)

 Table 1. Characteristics of participating cohorts.



Cohort name (country)	N	Birth years	Gestational age at birth (weeks)	Birth weight (gram)	FVC	FEV <sub>1</sub>	FEV₁/ FVC	FEF <sub>25-75</sub>	FEF <sub>75</sub>	Childhood asthma
			Median (5-95% range)	Mean (SD)	Mean Z-score (SD)	Mean Z-score (SD)	Mean Z-score (SD)	Mean Z-score (SD)	Mean Z-score (SD)	Yes
INMA Sabadell (Spain)	408	2004- 2007	39.8 (1.3)	3,261 (404)	-0.47 (1.38)	-0.57 (1.30)	-0.08 (1.03)	-0.61 (1.00)	-0.25 (1.12)	0.7 (3)
INMA Valencia (Spain)	455	2003- 2005	39.6 (1.7)	3,227 (491)	0.30 (1.10)	0.30 (1.08)	-0.04 (0.95)	-0.13 (0.91)	-0.04 (0.90)	-
ISLE OF WIGHT (United Kingdom)	1,030	1989- 1990	39.9 (1.5)	3,411 (510)	0.24 (0.91)	0.39 (1.01)	0.22 (1.03)	0.04 (0.99)	-	21.5 (221)
KOALA (The Netherlands)	438	2000- 2003	40.0 (1.2)	3,552 (467)	0.15 (0.94)	-0.13 (0.95)	-0.55 (0.84)	-	-	8.0 (35)
LEICESTER 1990 (United Kingdom)	290	1985- 1990	39.0 (2.2)	3,373 (599)	-0.33 (1.11)	-0.38 (1.12)	-0.76 (0.90)	-0.62 (1.01)	-	37.2 (108)
LEICESTER 1998 (United Kingdom)	1,476	1993- 1997	39.2 (2.0)	3,314 (592)	-0.41 (1.04)	-0.39 (1.05)	0.01 (1.03)	-	0.05 (0.94)	36.4 (538)
MAS (Germany)	641	1990	40.0 (1.4)	3,414 (460)	-0.06 (0.97)	0.24 (1.00)	0.41 (1.00)	1.15 (0.14)	-	5.0 (32)
PIAMA (The Netherlands)	1,767	1996- 1997	39.9 (1.7)	3,526 (540)	0.04 (0.95)	0.07 (1.04)	-0.04 (1.01)	-1.67 (1.21)	-0.21 (0.95)	10.0 (176)
RHEA (Greece)	666	2007- 2008	38.1 (1.7)	3,175 (506)	-0.25 (1.09)	-0.33 (1.14)	-0.10 (0.94)	-0.38 (0.96)	-0.17 (1.05)	5.9 (39)
SEATON (United Kingdom)	578	1997	39.5 (1.8)	3,488 (563)	-0.12 (1.08)	-0.06 (1.08)	-0.04 (0.96)	-0.27 (0.98)	-	20.1 (116)
SWS (United Kingdom)	803	1998- 2007	39.7 (1.9)	3,447 (548)	0.13 (1.01)	0.03 (0.95)	-0.18 (1.05)	-0.28 (0.94)	-	15.1 (121)
WHISTLER (The Netherlands)	591	2001- 2012	40.0 (1.3)	3,553 (499)	0.16 (1.11)	0.46 (1.14)	0.31 (0.93)	-0.04 (1.23)	0.12 (1.07)	9.3 (55)

 Table 1 (continued). Characteristics of participating cohorts.

N = number of participants with information on at least gestational age or birth weight, and a lung function outcome. Lung function outcomes are forced vital capacity (FVC), force expiratory volume in 1 second (FEV<sub>1</sub>), mid forced expiratory flow (FEF<sub>25-75</sub>) and force expiratory flow at 75% of the exhaled FVC (FEF<sub>75</sub>). Values are means (standard deviations) and percentages (absolute numbers) for the information on asthma. Additional information on data collection (Table S1), determinants (Table S2), outcomes (Table S3), and maternal and child related covariates (Tables S4, S5) is provided in the Supporting Information.

Table 2. Associations of birth weight, gestational age and infant weight gain with childhood asthma, additionally adjusted for lung function.

	<u>Risk of childhood asthma</u> Odds ratio (95% Confidence Interval)								
	Full model	<u>Full model</u> <u>+ FEV<sub>1</sub></u>	<u>% change</u> (95% Cl)	<u>Full model</u> <u>+ FEV₁/FVC</u>	<u>% change</u> (95% CI)	<u>Full model</u> + FEF <sub>75</sub>	<u>% change</u> (95% CI)		
<u>Gestational age (weeks)</u>	<u>0.94</u> (0.92, 0.97)** n = 15,019	<u>0.95</u> (0.93, 0.97)** n = 14,832	<u>-9.8%</u> (- 16.4, - 5.3)**	<u>0.95</u> (0.93, 0.97)** n = 14,017	<u>-13.5%</u> (-21.0, -7.3)**	<u>0.97</u> (0.94, 1.00) n = 9,177	<u>-44.6%</u> (-81.1, -14.6)**		
Preterm birth (<37 weeks)	<u>1.34</u> (1.15, 1.57)** n = 15,019	<u>1.30</u> (1.11, 1.53)** n = 14,832	<u>-7.3%</u> (-18.8, -0.9)*	<u>1.27</u> (1.08, 1.49)** n = 14,017	<u>-14.4%</u> (-39.6, -2.8)*	<u>1.20</u> (0.99, 1.47) <u>n = 9,177</u>	<u>-39.0%</u> (-69.3, -3.4)*		
<u>Birth weight (500 grams)</u>	<u>0.94</u> (0.90, 0.97)** n = 15,547	<u>0.95</u> (0.91, 0.99)* n = 15,360	<u>-18.9%</u> (-37.0, -11.2)**	<u>0.94</u> (0.90, 0.98)** n = 13,985	<u>-10.5%</u> (-21.9, -3.4)**	<u>0.96</u> (0.92, 1.02) <u>n = 9,135</u>	<u>-17.8</u> (-50.6, -9.0)**		
<u>Low birth weight (&lt;2,500 grams)</u>	<u>1.32</u> (1.07, 1.62)** n = 15,547	<u>1.25</u> (1.02, 1.54)* n = 15,360	<u>-19.0%</u> (-37.3, -11.8)**	<u>1.23</u> (0.99, 1.52) n = 13,985	<u>-21.6%</u> (-47.3, -11.4)**	<u>1.05</u> (0.81, 1.36) <u>n = 9,135</u>	<u>-82.5%</u> (-149, 10.3)		
<u>Birth weight (SDS)</u>	<u>0.98</u> (0.94, 1.03) n = 14,947	<u>1.00</u> (0.96, 1.05) n = 14,760	<u>-83.8%</u> (-950, 825)	<u>0.98</u> (0.94, 1.03) n =13,946	<u>-14.0%</u> (-247, 281)	<u>0.99</u> (0.93, 1.04) <u>n = 9,122</u>	<u>-15.8%</u> (-158, 169)		
Small for gestational age (<10th percentile)	<u>1.18</u> (1.01, 1.37)* n = 14,947	<u>1.13</u> (0.97, 1.32) <u>n = 14,760</u>	<u>-28.9%</u> (-253, 108)	<u>1.16</u> (0.99, 1.36) <u>n = 13,946</u>	<u>-18.8%</u> <u>(-123, 164)</u>	<u>1.20</u> (1.00, 1.44) <u>n = 9,122</u>	<u>10.2%</u> (-8.3, 26.2)		
Infant weight gain in first year (SDS), adjusted for gestational age and weight at birth	<u>1.27</u> (1.21, 1.34)** n = 12,511	<u>1.28</u> (1.22, 1.35)** n = 12,511	<u>6.5%</u> (2.3, 9.9)**	<u>1.25</u> (1.18, 1.31)** n = 11,780	<u>-8.4%</u> (-16.1, -3.2)**	<u>1.13</u> (1.06, 1.20)** <u>n = 7,969</u>	<u>-60.8</u> (-115, 39.5)		

\*p<0.05 \*\*p<0.01. Values are odds ratios or percentage change in odds ratios (95% confidence interval) from random effect models and represent the risk of asthma per week, 500 grams or SDS increase in gestational age, birth weight, gestational age adjusted birth weight (birth weight SDS), or infant weight gain (SDS), respectively, or represent odds ratios or percentage change in odds ratios (95% confidence interval) in risk of asthma for preterm birth vs. term birth, low birth weight vs. normal birth weight or small for gestational age vs. normal and large for gestational age (<10<sup>th</sup> percentile vs >10<sup>th</sup> percentile). Percentage change in odds ratio (OR) is calculated using the formula (100 x (OR<sub>mediator</sub> - OR<sub>model 1</sub>)/(OR<sub>model 1</sub> - 1)), with corresponding 95% confidence interval obtained by bootstrap procedures. To enable comparison of effect estimates, results for gestational age adjusted birth weight and infant weight gain are presented as per SDS. Models are adjusted for maternal education, smoking during pregnancy, smoking during childhood, atopy, asthma and child's sex, number of siblings, daycare attendance, breastfeeding, respiratory tract infections, childhood eczema, inhalant allergies and body mass index (full model), and additionally for lung function outcomes (mediator model). Table 2. Associations of birth weight, gestational age and infant weight gain with childhood asthma, additionally adjusted for lung function.

	Risk of childhood asthma Odds ratio (95% Confidence Interval)									
	Full model	<del>Full model</del> + FEV₁	<del>% change</del> <del>(95% Cl)</del>	<del>Full model</del> + FEV₁/FVC	<del>% change</del> <del>(95% Cl)</del>	Full model + FEF <sub>75</sub>	<del>% change</del> <del>(95% Cl)</del>			
Gestational age (weeks)	<del>0.94</del> <del>(0.92, 0.97)**</del>	<del>0.95</del> <del>(0.93, 0.97)**</del>	<del>-9.8%</del> <del>(- 16.4, - 5.3)**</del>	<del>0.95</del> <del>(0.93, 0.97)**</del>	<del>-13.5%</del> <del>(-21.0, -7.3)**</del>	<del>0.97</del> <del>(0.94, 1.00)</del>	<del>-44.6%</del> <del>(-81.1, -14.6)**</del>			
Preterm birth (<37 weeks)	1.34 (1.15, 1.57)**	<del>1.30</del> (1.11, 1.53)**	<del>-7.3%</del> <del>(-18.8, -0.9)*</del>	<del>1.27</del> <del>(1.08, 1.49)**</del>	<del>-14.4%</del> <del>(-39.6, -2.8)*</del>	<del>1.20</del> (0.99, 1.47)	<del>-39.0%</del> <del>(-69.3, -3.4)*</del>			
Birth weight (500 grams)	<del>0.94</del> <del>(0.90, 0.97)**</del>	0.95 (0.91, 0.99)*	<del>-18.9%</del> <del>(-37.0, -11.2)**</del>	0.94 (0.90, 0.98)**	<del>-10.5%</del> <del>(-21.9, -3.4)**</del>	<del>0.96</del> ( <del>0.92, 1.02)</del>	<del>-17.8</del> <del>(-50.6, -9.0)**</del>			
Low birth weight (<2,500 grams)	<del>1.32</del> <del>(1.07, 1.62)**</del>	(1.02, 1.54)*	- <del>19.0%</del> <del>(-37.3, -11.8)**</del>	<del>1.23</del> <del>(0.99, 1.52)</del>	- <del>21.6%</del> (-47.3, -11.4)**	(0.02, 1.02) 1.05 (0.81, 1.36)	<del>-82.5%</del> <del>(-149, 10.3)</del>			
Birth weight (SDS)	0.98 (0.94, 1.03)	<del>1.00</del> <del>(0.96, 1.05)</del>	<del>-83.8%</del> <del>(-950, 825)</del>	0.98 (0.94, 1.03)	<del>-14.0%</del> <del>(-247, 281)</del>	<del>0.99</del> <del>(0.93, 1.04)</del>	<del>-15.8%</del> <del>(-158, 169)</del>			
Small for gestational age (<10th	<del>1.18</del>	1.13	<del>-28.9%</del>	<del>1.16</del>	<del>-18.8%</del>	<del>1.20</del>	<del>10.2%</del>			
percentile)	<del>(1.01, 1.37)*</del>	<del>(0.97, 1.32)</del>	<del>(-253, 108)</del>	<del>(0.99, 1.36)</del>	<del>(-123, 164)</del>	(1.00, 1.44)	<del>(-8.3, 26.2)</del>			
Infant weight gain in first year (SDS),	1.27	1.28	<del>6.5%</del>	1.25	-8.4%	1.13	-60.8			
adjusted for gestational age and weight at	<del>(1.21, 1.34)**</del>	<del>(1.22, 1.35)**</del>	<del>(2.3, 9.9)**</del>	<del>(1.18, 1.31)**</del>	<del>(-16.1, -3.2)**</del>	<del>(1.06, 1.20)**</del>	<del>(-115, 39.5</del>			

birth

\*p<0.05 \*\*p<0.01. Values are odds ratios or percentage change in odds ratios (95% confidence interval) from random effect models and represent the risk of asthma per week, 500 grams or SDS increase in gestational age, birth weight, gestational age adjusted birth weight (birth weight SDS), or infant weight gain (SDS), respectively, or represent odds ratios or percentage change in odds ratios (95% confidence interval) in risk of asthma for preterm birth vs. term birth, low birth weight vs. normal birth weight or small for gestational age vs. normal and large for gestational age (<10<sup>th</sup> percentile vs >10<sup>th</sup> percentile). Percentage change in odds ratio (OR) is calculated using the formula (100 x (OR<sub>mediater</sub> - OR<sub>model 1</sub>)/(OR<sub>model 1</sub> - 1)), with corresponding 95% confidence interval obtained by bootstrap procedures. To enable comparison of effect estimates, results for gestational age adjusted birth weight and infant weight gain are presented as per SDS. Models are adjusted for maternal education, smoking during pregnancy, smoking during childhood, atopy, asthma and child's sex, number of siblings,

daycare attendance, breastfeeding, respiratory tract infections, childhood eczema, inhalant allergies and body mass index (full model), and additionally for

lung function outcomes (mediator model).