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## Article

# Recent experience with cognates and interlingual homographs in one language affects subsequent processing in another language

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### **Keywords:**

lexical decision; cognates; interlingual homographs; language switching; word-meaning priming.

### **Abstract**

This experiment shows that recent experience in one language influences subsequent processing of the same word-forms in a different language. Dutch–English bilinguals read Dutch sentences containing Dutch–English cognates and interlingual homographs, which were presented again 16 minutes later in isolation in an English lexical decision task. Priming produced faster responses for the cognates but slower responses for the interlingual homographs. These results show that language switching can influence bilingual speakers at the level of individual words, and require models of bilingual word recognition (e.g. BIA<sup>+</sup>) to allow access to word meanings to be modulated by recent experience.

## Introduction

Numerous studies have shown that switching between languages can produce general reductions in bilingual individuals' language processing speed. For example, latencies in picture naming, lexical decision and number naming tasks are slower when the preceding trial involves a stimulus or response in a different language (Costa & Santesteban, 2004; Grainger & Beauvillain, 1987; Meuter & Allport, 1999), indicating persistent interference from the previous language after a language switch. Furthermore, several studies have shown that bilinguals are affected by within-language ambiguity: recognising that a word is a translation of another word takes more time for words with multiple translations (e.g. "sacred" or "holy" for "heilig") than for words with a unique translation (e.g. "frog" for "kikker"; Boada, Sánchez-Casas, Gavilán, García-Albea & Tokowicz, 2013; Eddington & Tokowicz, 2013; Laxén & Lavour, 2010). However, few studies have investigated language switching effects on ambiguous words at the individual word level. Such cross-language priming effects could potentially exist both for cognates, which share a meaning across languages (e.g., the Dutch–English word "film"), and interlingual homographs, which have unrelated meanings (e.g., "room", meaning "cream" in Dutch).

Although current models of bilingual processing like the Bilingual Interactive Activation plus (BIA<sup>+</sup>) model (Dijkstra & van Heuven, 2002) make no clear predictions about such cross-language priming effects because they contain no mechanisms by which experience can influence performance at relatively long timescales, predictions about the precise nature of word-specific language switching effects come from analogous ambiguity effects in monolingual listeners. Just as "room" is ambiguous to Dutch–English bilinguals, "ball" is ambiguous to monolingual English speakers. Rodd, Cutrin, Kirsch, Millar and Davis (2013) found that, 20 minutes after hearing ambiguous words within disambiguating sentence contexts (e.g., "she found the perfect dress for the school ball"), listeners were biased to retrieve the primed meaning compared to an unprimed baseline. This 'word–meaning priming' endures at longer delays than semantic priming (when a synonym was presented during the priming phase), suggesting that the effect reflects a strengthening of the connection between the word-form representation and the primed meaning, such that the primed meaning becomes more readily available (possibly at the expense of alternative unprimed meanings). If an interlingual homograph's different meanings behave in a similarly competitive manner, an encounter with its Dutch meaning should delay access to the unrelated English meaning even minutes later. In contrast, facilitatory cross-language priming would be expected for cognates that share their meaning(s) in both languages, as the appropriate form-to-meaning mapping would be strengthened during priming.

These expectations rely critically on previous studies suggesting that bilingual speakers have one integrated lexicon and that access to it is language non-selective (see Dijkstra, 2005; Dijkstra & van Heuven, 2002; for a comprehensive review). Specifically, these studies show that (in the absence of priming manipulations) lexical decision times are faster for cognates than control words, suggesting cognates share a single representation in the bilingual lexicon, while response times to interlingual homographs are usually slower, due to competition between the two interpretations during lexical access. To model this, in the BIA<sup>+</sup> model, cognates share both their orthographic and semantic nodes. Presentation of the cognate activates letter features and letters, which activate these nodes. Resonance between the orthographic and semantic representations results in faster recognition of cognates relative to non-cognates

(Peeters, Dijkstra & Grainger, 2013). In contrast, interlingual homographs correspond to two orthographic nodes (one for each language) connected to two semantic nodes. The two orthographic nodes laterally inhibit each other and, presumably because they are identical, this competition is stronger than that between non-cognates, resulting in slower reaction times in comparison.

Previous work on this topic has produced mixed results. Cristoffanini, Kirsner and Milech (1986) primed cognates in a Spanish or English naming task and then measured English lexical decision times after a 10-minute delay. Similarly, Gerard and Scarborough (1989) primed cognates and interlingual homographs in a Spanish or English lexical decision task and measured either Spanish or English lexical decision times. Both studies reported facilitative priming in the cross-language conditions for cognates. Gerard and Scarborough also reported a facilitative effect for the interlingual homographs. In contrast, Lalor and Kirsner (2001), also using lexical decision tasks for both priming and testing, found no significant cross-language priming for interlingual homographs. Crucially, because these latter two studies used the same task for priming and testing, the results may have been influenced by task-specific priming that arises because participants make the same (positive) lexical decision response during both priming and testing (see Horner & Henson, 2009). For the interlingual homographs, this task-related facilitation may have cancelled out the expected disruptive effect of language switching. For this reason, we used different tasks for priming and testing.

The current study used a modified version of the monolingual word-meaning priming paradigm (Rodd et al., 2013). In the priming phase, Dutch–English bilinguals read *Dutch* sentences that contained either a cognate, an interlingual homograph or the Dutch translation of an English semantic control word (to create a semantic priming control condition). After a filler task lasting approximately ten minutes, the impact of priming was measured using an *English* lexical decision (LD) task. The semantic prime control condition, in which only the word's meaning (and not its form) is repeated, should reveal whether any observed priming in the cognate condition reflects general semantic priming distinct from repetition of the specific word-form. If the latter were true, we would observe priming for the semantic controls. The three priming conditions were each compared to an unprimed baseline.

## **Methods**

### ***Participants***

Thirty-two London-based Dutch–English bilinguals were paid for their participation. Three participants with mean lexical decision times above 1950ms were excluded. The remaining 29 participants (11 male; mean age 28.4 years,  $SD=7.2$  years) had on average 19.5 years' experience with English (minimum 11 years) and had been living in London on average for 4.8 years (range: 1 month–23 years). Participants rated their English proficiency an average of 8.8 out of 10. Participants' subjective ratings of their use of Dutch and English in daily life for talking, listening, reading, writing and thinking revealed they used English more than Dutch for all activities ( $p<.001$ ), except thinking ( $p=.33$ ).

## **Materials**

Thirty identical cognates, 30 identical interlingual homographs and 30 semantic control words were selected from previous studies (see Supplementary Materials Online; approximately half from Dijkstra, Grainger & van Heuven, 1999; the rest from other studies). Interlingual homographs were defined as words that did not share any (part) of their meaning(s). Cognates shared (at least) their dominant meaning. To obtain Dutch primes for the English semantic controls, 13 Dutch–English bilinguals (4 male; mean age 21.2 years) with a minimum of 9 years' experience with English (mean=12.7 years) translated 59 English words to Dutch using an online questionnaire. The 30 chosen semantic controls had a mean translation agreement of 92.7% (minimum=62%). As they had already been used in previous research, the cognates and interlingual homographs were not pre-tested. The three word types were matched on length, SUBTLEX word frequency (Brysbaert & New, 2009; Keuleers, Brysbaert & New, 2010; for English and Dutch, respectively) and orthographic similarity (using OLD20; Yarkoni, Balota & Yap, 2008; Table 1).

A Dutch sentence frame was constructed for each word (Table 2, Supplementary Materials Online). Six additional practice sentences were created. To minimise any semantic ambiguity, target words were placed towards the end of the sentence. The 30 sentences in each condition were pseudorandomly divided into two sets, matched for all key variables, for use in the two versions (see Procedure). The probes assigned to the sentences for the semantic relatedness task were either very strongly related or completely unrelated to the sentence.

Table 1. Means (and standard deviations) for all key matching variables and sentence length. Orthographic neighbourhood refers to Yarkoni et al.'s (2008) OLD20; Word frequency refers to the SUBTLEX word frequency in occurrences per million (see Brysbaert & New, 2009, for English and Keuleers et al., 2010, for Dutch); log-transformed word frequency refers to the SUBTLEX log transformed raw word frequency ( $\log_{10}[\text{raw frequency}+1]$ ); LG10WF in the SUBTLEX databases).

	DUTCH CHARACTERISTICS				ENGLISH CHARACTERISTICS			
	word length	sentence length	orthographic neighbourhood	word frequency	log-transformed word frequency	orthographic neighbourhood	word frequency	log-transformed word frequency
<b>cognates</b>	3.93 (0.64)	9.73 (1.58)	1.37 (0.31)	27.29 (30.69)	2.85 (0.47)	1.47 (0.31)	31.45 (35.49)	2.96 (0.47)
<b>interlingual homographs</b>	3.97 (0.72)	9.67 (1.56)	1.27 (0.29)	26.23 (27.65)	2.83 (0.50)	1.34 (0.27)	29.41 (36.24)	2.93 (0.49)
<b>semantic controls</b>	3.93 (0.78)	9.70 (1.44)				1.55 (0.23)	30.84 (35.97)	2.98 (0.43)

Table 2. Examples of prime sentences for the semantic relatedness task. The lexical decision task targets are underlined.

	prime sentence (Dutch original)	prime sentence (English translation)	probe (Dutch original)	probe (English translation)	related?
<b>cognate</b>	Hij nam elke dag de <u>bus</u> naar school.	He took the <u>bus</u> to school every day.	jaloerie	jealousy	No
<b>interlingual homograph</b>	Alleen vrouwelijke bijen en wespen hebben een <u>angel</u> .	Only female bees and wasps have a <u>sting</u> .	zee	sea	No
<b>semantic control</b>	De schrijver zat achter zijn <u>bureau</u> te schrijven.	The writer was writing at his <u>desk</u> .	tafel	table	Yes

The 90 non-words comprised fifteen Dutch words (e.g., “bron”), included to ensure participants only responded “yes” specifically to English words and not on the basis of overall familiarity, and 75 similar-length English-sounding pseudohomophones from previous studies (Rastle, Harrington & Coltheart, 2002; Rodd, 2000; Rodd, Gaskell & Marslen-Wilson, 2002). Pseudohomophones (instead of regular non-words) were used to encourage relatively deep processing (Rodd et al., 2002).

### ***Design and procedure***

Half the words of each word type were primed (i.e., appeared during the priming phase) while half were unprimed (i.e., only occurred in the later test phase). Two versions of the experiment were created such that participants saw each experimental item only once but across participants items occurred in both the primed and unprimed conditions.

The experiment comprised three separate tasks: a Dutch semantic relatedness task (approximately 5 minutes), an English digit span task (approximately 10 minutes) and an English lexical decision task (approximately 7 minutes). On average, LDs to words were made 16 minutes after they were primed. The three tasks were presented separately (using Matlab R2012a), with no indication that they were linked. Responses were recorded via a standard keyboard. At the end, participants completed a language background survey in Dutch.

#### ***Dutch semantic relatedness task***

This task served to prime the cognates, interlingual homographs and semantic controls. To ensure semantic processing, participants indicated via button presses whether the subsequent probe was semantically related to the preceding sentence. Participants read 3 practice sentences and 45 experimental sentences presented in different random orders. Each sentence remained on the screen for 4000ms; each probe remained on the screen until the participant responded. The inter-trial interval was 1000ms.

### *English digit span task*

This task introduced a delay, while minimising exposure to additional linguistic material. It was conducted in English to minimise any general language switch cost on the lexical decision task. Each string comprised 4-9 digits. Each digit was presented for 500ms with 500ms between trials. Participants saw 5 practice strings followed by 54 experimental strings divided into 3 blocks. 15-second breaks were enforced after each block.

### *English lexical decision task*

Participants saw all 180 stimuli (90 words, 45 of which were primed, and 90 non-words) and were asked to indicate via button presses, as quickly and accurately as possible, whether they were *real* English words or not. Twelve practice items were followed by 4 blocks of 45 stimuli. Block and item order was randomised across participants. Each block began with six fillers. Items remained on screen until participants responded, with a 500ms inter-trial interval.

## **Results**

### ***Semantic relatedness task***

High accuracy scores ( $M=93.8\%$  correct,  $SD=3.1\%$ , range 86.7%–97.8%) confirmed participants had processed the sentence meanings. There were no significant differences between the word types (cognates:  $M=95.0\%$ ,  $SD=10.6\%$ ; interlingual homographs:  $M=93.5\%$ ,  $SD=12.8\%$ ; semantic controls:  $M=92.9\%$ ,  $SD=22.1\%$ ) [ $F_1(2,62)=1.4$ ,  $p=.26$ ,  $\eta_p^2=.04$ ;  $F_2(2,87)=0.1$ ,  $p=.87$ ,  $\eta_p^2<.01$ ].

### ***Digit span***

Digit span (greatest string length recalled with at least 50% accuracy) was within normal limits ( $M=6.0$  digits, range 4–8 digits), confirming task engagement.

### ***Lexical decisions***

One cognate (“ark”) and two interlingual homographs (“arts”, “genie”) were excluded from analysis due to accuracy rates below 60%.<sup>1</sup> Reaction times (RTs) for incorrect trials were discarded (4.1% of the data), as were RTs more than two standard deviations above a participant’s mean (2.2% of the remaining data). All data points were above 300ms. Mean RTs and accuracy were calculated for each condition across participants and items.<sup>2</sup>

### *Reaction times*

Because the hypotheses concerned specific comparisons between two of the three word types, two  $2 \times 2 \times 2$  repeated measures ANOVAs were conducted with three factors: word type, priming manipulation and version (Figure 1A; see Rodd et al., 2013, Experiment 3,

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<sup>1</sup> Removing these words did not appreciably affect the matching of the stimuli.

<sup>2</sup> The analyses were repeated after inverse-transforming the data (Ratcliff, 1993) to examine the influence of remaining outliers; the significance levels were the same in both analyses.

for a similar approach). Main effects and interactions including version are not reported (Pollatsek & Well, 1995).

Comparing cognates and interlingual homographs revealed a significant main effect of word type [ $F_1(1,27)=10.2$ ,  $p=.004$ ,  $\eta_p^2=.27$ ;  $F_2(1,53)=10.3$ ,  $p=.002$ ,  $\eta_p^2=.16$ ], where cognates were recognised faster than interlingual homographs, but no effect of priming (both  $ps>.40$ ). Importantly, the word type  $\times$  priming interaction was significant [ $F_1(1,27)=5.6$ ,  $p=.03$ ,  $\eta_p^2=.17$ ;  $F_2(1,53)=5.3$ ,  $p=.03$ ,  $\eta_p^2=.09$ ]: priming had a facilitatory effect for cognates, but an inhibitory effect for interlingual homographs.

Comparing cognates and semantic controls revealed no main effect of word type (both  $ps>.10$ ). The effect of priming was marginally significant only in the participants analysis [ $F_1(1,27)=3.1$ ,  $p=.09$ ,  $\eta_p^2=.10$ ;  $F_2(1,55)=2.5$ ,  $p=.12$ ,  $\eta_p^2=.04$ ], such that primed words were recognised faster than unprimed words. There was no significant interaction between word type and priming (both  $ps>.10$ ).

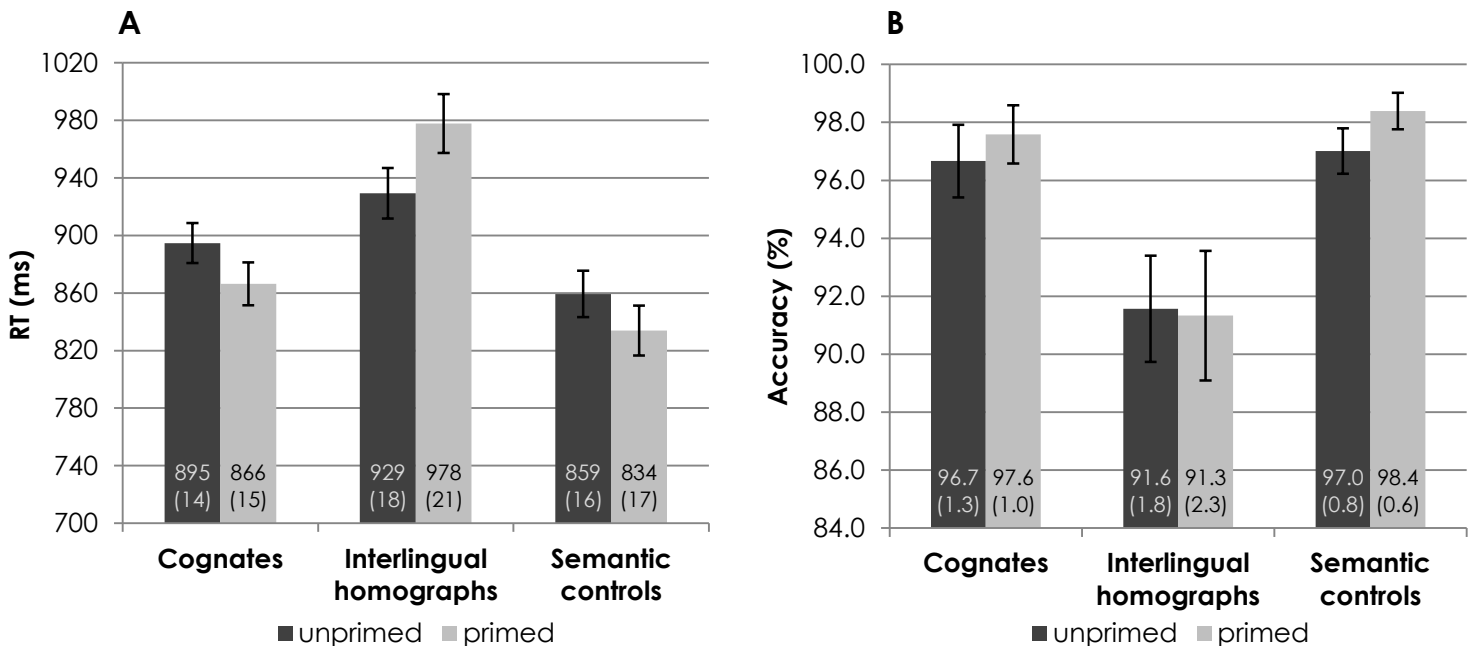


Figure 1. (A) Means of the lexical decision reaction times (in milliseconds) for the participants analysis by word type (cognates, interlingual homographs and semantic controls) and priming (unprimed and primed). Error bars represent the standard error of the mean adjusted for a within-participants design (Loftus & Masson, 1994). (B) Means of the lexical decision accuracy data (in per cent correct) for the participants analysis by word type (cognates, interlingual homographs and semantic controls) and priming (unprimed and primed). Error bars represent the standard error of the mean adjusted for a within-participants design (Loftus & Masson, 1994).

A further three  $2 \times 2$  repeated measures ANOVAs with factors version and priming were conducted to determine whether priming had been effective for each word type separately. These revealed a marginally significant facilitative effect of priming by participants for the cognates [ $F_1(1,27)=4.1$ ,  $p=.054$ ,  $\eta_p^2=.15$ ;  $F_2(1,27)=1.95$ ,  $p=.17$ ,  $\eta_p^2=.07$ ], but a marginally significant disruptive effect for the interlingual homographs



$[F_1(1,27)=3.6, p=.07, \eta_p^2=.12; F_2(1,26)=3.3, p=.08, \eta_p^2=.11]$ . The effect of priming was not significant for the semantic controls (both  $ps>.25$ ).

#### *Accuracy data*

Comparing cognates and interlingual homographs revealed a significant main effect of word type [ $F_1(1,27)=21.6, p<.001, \eta_p^2=.44; F_2(1,53)=10.0, p<.001, \eta_p^2=.16$ ], where cognates were processed more accurately than interlingual homographs. No other effects were significant (all  $ps>.40$ ; Figure 1B).

Comparing cognates and semantic controls revealed a marginally significant main effect of priming by participants only [ $F_1(1,27)=3.2, p=.09, \eta_p^2=.11; F_2(1,55)=2.6, p=.11, \eta_p^2=.05$ ], where primed items were processed more accurately than unprimed items. No other effects were significant (all  $ps>.60$ ).

### **General discussion**

This experiment shows that a single encounter with a cognate or interlingual homograph in one language can affect its subsequent processing in another language after an average delay of 16 minutes, and that this priming effect is influenced by the relationship between the Dutch and the English meanings. Before discussing these findings in detail, it should be noted that although overall lexical decision task performance may be influenced by a general language switching effect (which we attempted to minimise by conducting the preceding digit span task in English), any such effect would influence all items including the unprimed controls, and, therefore, cannot contribute to the observed priming effects.

Specifically, the cognates showed a 28ms facilitatory priming effect, consistent with previous studies (Cristoffanini et al., 1986; Gerard & Scarborough, 1989). In contrast, the interlingual homographs showed a 49ms inhibitory effect. This significant interaction between word type and priming seems inconsistent with the results of Gerard and Scarborough (1989) and Lalor and Kirsner (2001), who found either facilitation or no priming with cross-language repetition of interlingual homographs. We suggest that in these studies, use of the same task for priming and testing phases resulted in masking of the interference effect by facilitatory priming due to stimulus-response binding, as in both tasks each critical item would be mapped to the same “yes” response (cf. Horner & Henson, 2009), thereby masking the interference effect.

These results are consistent with the proposed explanation of word-meaning priming in the monolingual domain (Rodd et al., 2013): exposure to a word strengthens the connection between its form and the contextually appropriate meaning, so that during subsequent encounters with that word the primed meaning is more readily available, while access to the unprimed meaning is disrupted. This experiment suggests that similar processes operate in the bilingual domain: strengthening the form-to-meaning mapping enhances performance for words which share their form and meaning across languages (cognates), but interferes with processing of words that share their form but not their meaning (interlingual homographs). These findings are also consistent with claims that bilingual speakers’ lexical representations are not accessed in a language-specific manner (see Dijkstra, 2005; Dijkstra & van Heuven, 2002).

There was an unexpected non-significant 25ms ‘priming effect’ for the semantic control words, such that priming for semantic controls and cognates did not differ significantly, a pattern inconsistent with the absence of semantic priming in Rodd et al. (2013, Exp. 3). Most likely, however, this priming was not semantic in nature, but reflects long-term cross-language repetition (or translation) priming, such that upon presentation of the Dutch translation of the English semantic control, the semantic control word itself was also accessed and the connection between its form and meaning strengthened. Consistent with this account, Zeelenberg and Pecher (2003) found long-term cross-language repetition priming with high-fluency Dutch–English bilinguals using non-cognates, though only when both the priming and testing task were conceptual in nature. Li, Mo, Wang, Luo and Chen (2009) obtained similar results with low-fluency Chinese–English bilinguals. As there are some differences between the current study and those of Zeelenberg and Pecher (2003) and Li et al. (2009), most notably in the direction of the priming effect (they found L2-to-L1 priming only, whereas we observed L1-to-L2 priming) and the use of different tasks, further research is needed to determine whether the non-significant priming effect observed here was semantic in nature or reflects long-term translation priming.

In its current form, the BIA<sup>+</sup> model (Dijkstra & van Heuven, 2002) cannot explain these results. The BIA<sup>+</sup> framework incorporates *short-term* (cross-language) priming, such that each encounter with a word elevates its resting activation level (and decreases that of its neighbours). On subsequent encounters, this increase in nodal resting activity means that the recognition threshold is reached more quickly for regular words and cognates, and more slowly for interlingual homographs. However, this increase in resting activity is transient and is, therefore, unlikely to underlie the longer-term priming effect observed here. In essence, the BIA<sup>+</sup> is restricted to recognition and decision processes and is not a model that can learn from experience. As in the original Interactive Activation model (McClelland & Rumelhart, 1981), all parameters in the model are set by hand (Dijkstra, Hilberink-Schulpen & van Heuven, 2010) rather than being learned on the basis of experience with the lexical items. Consequently, the BIA<sup>+</sup> model, in its current form, cannot directly explain or predict how recognition processes are affected by long-term language switching, a situation bilinguals encounter on a regular basis (but see Grainger, Midgley & Holcomb, 2010, for suggested extensions of the model to include learning). Although connectionist models that include learning algorithms could potentially accommodate the current findings, such a model (of long-term word-meaning priming) has not yet been implemented (see Rodd et al., 2013 for more).

Finally, the current experiment also contributes to the monolingual word recognition literature. In the monolingual domain, Rodd et al. (2013) showed that long-term word-meaning priming can alter people’s meaning preferences, as revealed by unsped responses on a word association task. This experiment confirms the presence of word-meaning priming on a speeded lexical decision task that is less susceptible to potential effects of demand characteristics than unsped word association. Together with previous monolingual studies, the current experiment supports the view that the interpretation of ambiguous words is strongly influenced by recent experience: meaning preferences are not stable, but are a fluid and dynamic, property of a mental lexicon that is constantly changing in response to experience.

## References

- Boada, R., Sánchez-Casas, R., Gavilán, J. M., García-Albea, J. E., & Tokowicz, N. (2013). Effect of multiple translations and cognate status on translation recognition performance of balanced bilinguals. *Bilingualism: Language and Cognition*, *16*, 183-197.
- Brysbaert, M., & New, B. (2009). Moving beyond Kučera and Francis: A critical evaluation of current word frequency norms and the introduction of a new and improved word frequency measure for American English. *Behavior Research Methods*, *41*, 977-990.
- Costa, A., & Santesteban, M. (2004). Lexical access in bilingual speech production: Evidence from language switching in highly proficient bilinguals and L2 learners. *Journal of Memory and Language*, *50*, 491-511.
- Cristoffanini, P., Kirsner, K., & Milech, D. (1986). Bilingual lexical representation: The status of Spanish-English cognates. *The Quarterly Journal of Experimental Psychology*, *38*, 367-393.
- Dijkstra, T. (2005). Bilingual visual word recognition and lexical access. In J. F. Kroll & A. M. B. De Groot (eds.), *Handbook of bilingualism: Psycholinguistic approaches*, pp. 179-201. New York, NY: Oxford University Press.
- Dijkstra, T., Grainger, J., & van Heuven, W. J. B. (1999). Recognition of cognates and interlingual homographs: The neglected role of phonology. *Journal of Memory and Language*, *41*, 496-518.
- Dijkstra, T., Hilberink-Schulpen, B., & van Heuven, W. J. B. (2010). Repetition and masked form priming within and between languages using word and nonword neighbors. *Bilingualism: Language and Cognition*, *13*, 341-357.
- Dijkstra, T., & van Heuven, W. J. B. (2002). The architecture of the bilingual word recognition system: From identification to decision. *Bilingualism: Language and Cognition*, *5*, 175-197.
- Eddington, C. M., & Tokowicz, N. (2013). Examining English-German translation ambiguity using primed translation recognition. *Bilingualism: Language and Cognition*, *16*, 442-457.
- Gerard, L. D., & Scarborough, D. L. (1989). Language-specific lexical access of homographs by bilinguals. *Journal of Experimental Psychology: Learning, Memory, and Cognition*, *15*, 305-315.
- Grainger, J., & Beauvillain, C. (1987). Language blocking and lexical access in bilinguals. *The Quarterly Journal of Experimental Psychology*, *39*, 295-319.
- Grainger, J., Midgley, K., & Holcomb, P. J. (2010). Re-thinking the bilingual interactive-activation model from a developmental perspective (BIA-d). In M. Kali & M. Hickmann (eds.), *Language acquisition across linguistic and cognitive systems* (Vol. 52), pp. 267-283. Amsterdam/Philadelphia: John Benjamins.
- Horner, A. J., & Henson, R. N. (2009). Bindings between stimuli and multiple response codes dominate long-lag repetition priming in speeded classification tasks. *Journal of Experimental Psychology: Learning, Memory, and Cognition*, *35*, 757-779.
- Keuleers, E., Brysbaert, M., & New, B. (2010). SUBTLEX-NL: A new measure for Dutch word frequency based on film subtitles. *Behavior Research Methods*, *42*, 643-650.
- Lalor, E., & Kirsner, K. (2001). The representation of "false cognates" in the bilingual lexicon. *Psychonomic Bulletin & Review*, *8*, 552-559.
- Laxén, J., & Lavaur, J.-M. (2010). The role of semantics in translation recognition: Effects of number of translations, dominance of translations and semantic relatedness of multiple translations. *Bilingualism: Language and Cognition*, *13*, 157-183.

- Li, L., Mo, L., Wang, R., Luo, X., & Chen, Z. (2009). Evidence for long-term cross-language repetition priming in low fluency Chinese–English bilinguals. *Bilingualism: Language and Cognition*, *12*, 13–21.
- Loftus, G. R., & Masson, M. E. J. (1994). Using confidence intervals in within-subject designs. *Psychonomic Bulletin & Review*, *1*, 476–490.
- McClelland, J. L., & Rumelhart, D. E. (1981). An interactive activation model of context effects in letter perception, part i: An account of basic findings. *Psychological review*, *88*, 375–407.
- Meuter, R. F. I., & Allport, A. (1999). Bilingual language switching in naming: Asymmetrical costs of language selection. *Journal of Memory and Language*, *40*, 25–40.
- Peeters, D., Dijkstra, T., & Grainger, J. (2013). The representation and processing of identical cognates by late bilinguals: RT and ERP effects. *Journal of Memory and Language*, *68*, 315–332.
- Pollatsek, A., & Well, A. D. (1995). On the use of counterbalanced designs in cognitive research: A suggestion for a better and more powerful analysis. *Journal of Experimental Psychology: Learning, Memory, and Cognition*, *21*, 785–794.
- Rastle, K., Harrington, J., & Coltheart, M. (2002). 358,534 nonwords: The ARC nonword database. *Quarterly Journal of Experimental Psychology, Section A: Human Experimental Psychology*, *55*, 1339–1362.
- Ratcliff, R. (1993). Methods for dealing with reaction time outliers. *Psychological Bulletin*, *114*, 510–532.
- Rodd, J. M. (2000). *Semantic representation and lexical competition: Evidence from ambiguity*. (Ph.D. dissertation, Cambridge University.).
- Rodd, J. M., Cutrin, B. L., Kirsch, H., Millar, A., & Davis, M. H. (2013). Long-term priming of the meanings of ambiguous words. *Journal of Memory and Language*, *68*, 180–198.
- Rodd, J. M., Gaskell, G., & Marslen-Wilson, W. (2002). Making sense of semantic ambiguity: Semantic competition in lexical access. *Journal of Memory and Language*, *46*, 245–266.
- Yarkoni, T., Balota, D., & Yap, M. (2008). Moving beyond Coltheart's N: A new measure of orthographic similarity. *Psychonomic Bulletin & Review*, *15*, 971–979.
- Zeelenberg, R., & Pecher, D. (2003). Evidence for long-term cross-language repetition priming in conceptual implicit memory tasks. *Journal of Memory and Language*, *49*, 80–94.