

**Identification of common genetic variants
influencing spontaneous dizygotic twinning
and female fertility**

Hamdi Mbarek,^{1,21,*} Stacy Steinberg,^{2,22} Dale R. Nyholt,^{3,4,22} Scott D. Gordon,⁴ Michael B. Miller,⁵ Allan F. McRae,^{6,4} Jouke Jan Hottenga,^{1,21} Felix R. Day,²⁰ Gonneke Willemsen,^{1,21} Eco J. de Geus,^{1,21} Gareth E. Davies,⁷ Hilary C. Martin,⁸ Brenda W. Penninx,⁹ Rick Jansen,⁹ Kerrie McAloney,⁴ Jacqueline M. Vink,¹ Jaakko Kaprio,¹⁰ Robert Plomin,¹¹ Tim D. Spector,¹² Patrik K. Magnusson,¹³ Bruno Reversade,¹⁴ R. Alan Harris,¹⁵ Kjersti Aagaard,¹⁵ Ragnar P. Kristjansson,² Isleifur Olafsson,¹⁶ Gudmundur Ingi Eyjolfsson,¹⁷ Olof Sigurdardottir,¹⁸ William G. Iacono,⁵ Cornelis B. Lambalk,¹⁹ Grant W. Montgomery,⁴ Matt McGue,⁵ Ken K. Ong,²⁰ John R.B Perry,²⁰ Nicholas G. Martin,⁴ Hreinn Stefánsson,² Kari Stefánsson,² and Dorret I. Boomsma,^{1,21,*}

¹Department of Biological Psychology, [Vrije U-Amsterdamniversiteit](https://www.vrijeuniversiteit.nl), 1081BT Amsterdam, The Netherlands.

²deCODE Genetics, 101 Reykjavik, Iceland.

³Institute of Health and Biomedical Innovation, Queensland University of Technology, 4059 Brisbane, Australia.

⁴QIMR Berghofer Medical Research Institute, 4029 Brisbane, Australia.

⁵Department of Psychology, University of Minnesota, Minneapolis, MN 55455, USA.

⁶Queensland Brain Institute, The University of Queensland, 4072 Brisbane, Australia.

⁷Avera Institute for Human Genetics, Sioux Falls, SD 57108, USA.

⁸Wellcome Trust Centre for Human Genetics, OX3 TBN Oxford, United Kingdom.

⁹Department of Psychiatry and EMGO Institute for Health and Care Research, [Vrije U Universiteit](https://www.vrijeuniversiteit.nl), Medical Center/GGZ inGeest, 1081HL Amsterdam, The Netherlands.

¹⁰Institute for Molecular Medicine Finland FIMM, University of Helsinki; Department of Public Health, University of Helsinki; National Institute for Health and Welfare, FI-00014 Helsinki, Finland.

¹¹Medical Research Council Social, Genetic and Developmental Psychiatry Centre, Institute of Psychiatry, Psychology & Neuroscience, King's College London, SE5 8AF London, United Kingdom.

¹²Department of Twin Research & Genetic Epidemiology, King's College London, SE1 7EH London, United Kingdom.

¹³Department of Medical Epidemiology and Biostatistics, Karolinska Institute, SE-171 77 Stockholm, Sweden.

¹⁴Institute of Medical Biology, A*STAR, 138648 Singapore, Singapore.

¹⁵Division of Maternal-Fetal Medicine, Departments of Obstetrics and Gynecology, Molecular and Human Genetics, and Molecular and Cell Biology, Baylor College of Medicine and Texas Children's Hospital, Houston, TX 77030, USA.

¹⁶Department of Clinical Biochemistry, Landspítali- The National University Hospital of Iceland, 101 Reykjavik, Iceland.

¹⁷Icelandic Medical Center (Laeknasetrid), Laboratory in Mjodd (RAM), Reykjavik, Iceland.

¹⁸Department of Clinical Biochemistry, Akureyri Hospital, 109 Akureyri, Iceland.

¹⁹Department of Obstetrics, Gynecology and Reproductive Medicine, VU University Medical Center, 1007MB Amsterdam, The Netherlands.

²⁰MRC Epidemiology Unit, University of Cambridge School of Clinical Medicine, Institute of Metabolic Science, Cambridge Biomedical Campus, CB2 0QQ Cambridge, United Kingdom.

²¹EMGO+ Institute for Health and Care Research, 1081BT Amsterdam, The Netherlands.

²²These authors contributed equally to this study

*Correspondence to h.mbarek@vu.nl (H.M, @drHamdiMbarek), di.boomsma@vu.nl (D.I.B)

Abstract

Spontaneous dizygotic (DZ) twinning occurs in 1-4% of women, with familial clustering and unknown ~~patho~~physiological pathways and genetic origin. DZ twinning may index increased fertility and has distinct health implications for mother and child. We performed ~~the first a~~ GWA study in 1,980 mothers of spontaneous DZ twins and 12,953 controls. Findings were replicated in a large Icelandic cohort and tested for association across a broad range of fertility traits in women. Two SNPs were identified (rs11031006 near *FSHB*, $p = 1.54 \times 10^{-9}$, and rs17293443 in *SMAD3*, $p = 1.57 \times 10^{-8}$) and replicated ($p = 3 \times 10^{-3}$ and $p = 1.44 \times 10^{-4}$, respectively). Based on ~90,000 births in Iceland, the relative risk of a mother delivering twins increased by 18% for each copy of ~~the allele~~ rs11031006-G, and 9% for ~~the~~ rs17293443-C. A higher polygenic risk score (PRS) for DZ twinning, calculated based on the results of the DZ twinning GWAS, was significantly associated with DZ twinning in Iceland ($p = 0.001$). A higher PRS was also associated with having children ($p = 0.01$), greater lifetime parity ($p = 0.03$) and earlier age at first child ($p = 0.02$). ~~The Allele~~ rs11031006-G was associated with higher serum FSH levels, earlier age at menarche, earlier age at first child, higher lifetime parity, lower PCOS risk, and earlier age at menopause. Conversely, ~~the~~ rs17293443-C was associated with later age at last child. We identified ~~the first~~ robust genetic risk variants for DZ twinning: one near *FSHB*, and a second within *SMAD3*, ~~—~~the product of which plays an important role in gonadal responsiveness to FSH. These loci contribute to crucial aspects of reproductive capacity and health.

1 **Introduction**

2 DZ twinning (MIM: 276400) is defined as the concomitant conception and
3 development of two independent zygotes during one pregnancy. Mothers of spontaneous
4 (conceived without assisted reproductive technology) DZ twins have a predisposition to
5 multiple ovulation events due to interference with single dominant follicle selection, a
6 biological mechanism fundamental for the human species.¹ DZ twinning is common, with
7 large regional differences from six per 1000 births in Asia to 40 per 1000 births in Africa,²
8 whereas monozygotic twinning occurs around the world at a constant rate of around three to
9 four per 1000 births.^{3; 4} DZ twinning rates also vary substantially over time. For example, in
10 the US, the observed incidence of twin births increased by a factor of 1.9 between 1971 and
11 2009.⁵ While a considerable part of the increase is attributable to fertility treatments, with an
12 estimate of 36% of all twins born in the USA in 2011 resulting from assisted reproduction, the
13 majority of twins are still conceived spontaneously. In addition to fertility treatments,
14 increases in maternal age contribute to increases in twinning incidence.^{3; 6}

15 Twinning is associated with increased risks to mother and offspring, including higher
16 risks of stillbirth, neonatal death and premature birth. Compared to singleton children, twins
17 use more hospital resources, especially during the first year of life,⁷ with hospital costs in the
18 first five years of life being as much as 3-3-fold higher.⁸

19 Mothers of DZ twins differ from other women in that they are taller, have an increased
20 BMI, are more often overweight and more often smoke before the twin pregnancy.⁹ Family
21 history, increased parity and gravidity all increase the risk of spontaneous DZ twinning.^{1; 3}
22 Remarkably, twinning rates do not reflect average nutritional status of a population, as
23 established from longitudinal studies in countries that experienced periods of starvation, such
24 as the Dutch hunger winter. Above a specific, yet undetermined, threshold, nutrition seems to
25 be of minimal importance for reproduction in general and also for twinning.¹⁰ These

26 observations all point to spontaneous twinning being a heritable trait and suggest the potential
27 for polygenic inheritance. In a landmark study of twinning based on data from genealogic
28 records from Utah,¹¹ White and Wyshak established that the genotype of the mother, but not
29 that of the father, affects the frequency of DZ twinning.

30 The underlying physiological mechanism for DZ twinning is the release and
31 fertilization of two or more oocytes. Ovarian folliculogenesis and determination of ovulation
32 quota are controlled both by circulating concentrations of follicle-stimulating hormone (FSH)
33 and by intra-ovarian factors including the two oocyte growth factors, *GDF9* (MIM: 601918)
34 and *BMP15* (MIM: 300247), as well as their cognate receptors.¹² In the common marmoset
35 monkey (subfamily Callitrichinae), DZ twins comprise the predominant litter size, and
36 singletons are rarely, if ever, observed. On DNA sequencing, specific nonsynonymous
37 substitutions were identified in *GDF9*, *BMP15*, *BMP4* (MIM: 112262), and *WFIKKN1* (MIM:
38 608021) as having a role in Callitrichine twinning.¹³ These genes are among a larger set of 63
39 candidate genes with a potential involvement in regulation of ovulation number and/or control
40 of growth and body size.¹⁴

41 Efforts to characterize the genes that contribute to DZ twinning in humans have not
42 been successful. Candidate gene and genome-wide linkage studies failed to uncover common
43 variants associated with DZ twinning,¹⁵⁻¹⁹ although one study reported rare variants in *GDF9*
44 to be associated with DZ twinning.¹⁸

45 DZ twinning has been suggested as a measure of human fertility both at the individual
46 and at the population level.²⁰ Spontaneous DZ twinning may be considered as a marker of
47 high fertility, as it reflects the frequency of double ovulation, the probability of coitus within
48 the appropriate time frame with fertilization of both ova, and maintenance of a multiple
49 pregnancy.

50 | The aim of this study was to perform ~~the first a~~ genome-wide association study
51 | [\(GWAS\)](#) in mothers of spontaneous DZ twins to identify relevant genomic regions and test
52 | their significance across a broad range of female fertility and reproductive traits including age
53 | at menarche, age at natural menopause, age at first and last child, and lifetime parity. [Three](#)
54 | [twin registries, from the Netherlands, Australia and Minnesota \(USA\) had detailed](#)
55 | [information on spontaneous twinning in mothers of DZ twins, as well as genotype data.](#)
56 | [Replication of top hits for twinning was possible in the Icelandic population and for other](#)
57 | [measures of reproductive ageing in several large-scale population meta-analyses](#)²¹⁻²⁴ .
58 |

59 | **Material and Methods**

60 | ~~Details of methods and associated references can be found in the supplemental data.~~

61 | **Descriptions of Participating Studies**

62 | *Netherlands Twin Register (NTR)*

63 | [The NTR sample consisted of 806 cases and 4,535 controls from the Netherlands Twin](#)
64 | [Register \(2,776 participants\) and the Netherlands Study of Depression and anxiety \(NESDA;](#)
65 | [2,565 participants\). NTR participants were ascertained by the presence of liveborn twins or](#)
66 | [triplets in the family and consist of multiples, their parents, siblings and spouses. Twins were](#)
67 | [born in all strata of society and NTR represents a general sample from the Dutch population](#)²⁵⁻
68 | ²⁸ [. NESDA is a longitudinal study focusing on the course and consequences of depression and](#)
69 | [anxiety disorders. Subjects for NESDA were recruited from the general population, mental](#)
70 | [health organizations and general practices. The sample includes subjects selected for](#)
71 | [depression and anxiety, as well as healthy controls](#)²⁹ . [Zygosity of twins was confirmed by](#)
72 | [DNA genotyping. Data on mode of pregnancy were available from several data collection](#)
73 | [waves including surveys sent out to mothers of twins, a survey to parents upon registration of](#)
74 | [young twins, and telephone interview as part of a project on DZ twinning](#)⁹ . The comparison of

75 the survey data with the hospital records showed that mothers can accurately report on the
76 mode of conception of their twins³⁰. Participants were excluded if they reported the use of
77 assisted reproductive technology at one or more occasions. In case no reports on mode of
78 pregnancy were available, data were excluded unless the twins were born prior to 1985.
79 *QIMR Berghofer Medical Research Institute (QIMR)*
80 The sample used in this analysis consisted of 606 cases and 6,656 controls. The individuals
81 were drawn from families containing (any type of) twins recruited for prior studies, either
82 from around Australia from the Australian Twin Registry (ATR) (generally twins born before
83 1971) or from south-eastern Queensland (generally twins born after 1980). Study recruitment
84 was predominantly population based (any family where the twins were willing to participate)
85 with no screening performed on reproductive phenotypes apart from selecting families with
86 twins. Studies for the older cohort typically were focused on personality traits but not selected
87 for them. Zygosity of twins was reported at time of recruitment; during phenotyping studies;
88 and tested by genotyping (on SNP arrays or Sequenom assays). New phenotyping excluded
89 mothers ('cases') from Queensland cohort, who used Assisted Reproductive Technology
90 (ART—typically IVF or hormone treatment) to become pregnant. Screening questions asked of
91 the mother during clinical sessions for phenotyping, were the primary basis for excluding
92 ART cases. A smaller subset of mothers was contacted specifically to establish this
93 information where it was not otherwise available. For the older (ATR) cohort, at the time of
94 the twins' birth, IVF was not yet in clinical use and other ART was rare.
95 *Minnesota Center for Twin and Family Research (MCTFR)*
96 All subjects in this sample were independently ascertained through vital records of the State
97 of Minnesota in an effort to construct a population-based twin registry^{31; 32}. The sample for
98 the current study consisted of 568 mothers of DZ twins and 1,862 controls who were the
99 parents of MZ twins from 1,062 families, including 800 complete parental pairs, 203 mothers

100 [and 59 fathers. Most of the twins were born in the 70s or early 80s, when even though fertility](#)
101 [treatment was available in the US, it was expensive and few had access to it. Genotyping was](#)
102 [population-based and independent of phenotypes other than twinning. About 92% of the](#)
103 [registry, and 100% of both case and control samples, are of primarily European ancestry.](#)
104 [Iceland \(deCODE\)](#)
105 [Mothers of twins or other multiples \(“twins”\) were selected from among those taking part in](#)
106 [deCODE’s genetic studies based on a nationwide genealogical database. To increase the](#)
107 [proportion of these mother of twins who were mothers of dizygotic twins, twins who had been](#)
108 [genotyped and shown to be monozygotic were not used to identify mothers. Controls were](#)
109 [individuals participating in deCODE’s genetic research from which both mothers of twins and](#)
110 [the mothers’ first-degree relatives had been removed. For the prediction of twinning using](#)
111 [polygenic risk scores, mothers having opposite sex or verified dizygotic twins were compared](#)
112 [with mothers who did not have twins.](#)

114 **Study Design, Genome-Wide Association Study and Replication**

115 We established the Twinning GWAS Consortium (TGC) to characterize the genetic basis of
116 DZ twinning in humans and performed [the first a](#) genome-wide association study (GWAS)
117 for DZ twinning utilizing [data from](#) 1,980 mothers of DZ twins (MODZT) and 12,953
118 controls from ~~The Netherlands, US and Australian~~ European ancestry cohorts. Sample sizes
119 and study characteristics are described in table [S1](#). All cases underwent screening to exclude
120 mothers who received assisted reproductive techniques. Controls were screened to exclude
121 pedigrees containing DZ twins. In the replication stage, significant findings from the meta-
122 analysis were tested in large Icelandic cohort of 3,597 mothers of twins and 297,348 controls.
123 All participants provided written informed consent, including consent for genotyping and

124 analysis, [and were recruited according to the protocols approved by the institution review](#)
125 [board of each institution.](#)

127 **Fertility Traits Measures**

128 Fertility measures (having children, number of children, age at first and last child and average
129 birth interval) were defined in females born before 1970 based on the Icelandic genealogy.
130 [Data for age at menarche, age at menopause and polycystic ovary syndrome were derived](#)
131 [from previously published GWAS consortia²¹⁻²⁴. Sample sizes and study characteristics are](#)
132 [described in table 2.](#)

134 **Genotyping, Quality Control and Imputation**

135 Each participating cohort performed participant-level genotyping of single nucleotide
136 polymorphisms (SNPs), that included standard quality-control measures for genotyping and
137 imputation (Table S21 [and Supplemental Methods](#)). All imputations were performed using the
138 1000 Genomes Project March 2012 release as the reference panel.

139 **Statistical analysis**

140 ~~GWA analysis was performed in each cohort using logistic regression under an additive~~
141 ~~genetic model with adjustment for principal components of genetic ancestry and specific~~
142 ~~covariates for each study. GWA results for each SNP (odds ratio [OR] and 95% confidence~~
143 ~~interval [CI]) were combined across cohorts using fixed-effect meta-analysis with inverse~~
144 ~~variance weighting. The most significant SNP at each genome-wide significant ($p < 5 \times 10^{-8}$)~~
145 ~~locus was tested in the replication stage and replicating SNPs were examined for association~~
146 ~~with FSH levels and other fertility traits.~~

147 **Association Tests**

148 [Each genome-wide association analysis from the three cohorts was conducted using logistic](#)
149 [regression under an additive genetic model with adjustment for principal components of](#)
150 [genetic ancestry. Because the GWAS data include family members we added the –family](#)
151 [option in the analysis, which takes the familial structure of the data into account using a](#)
152 [sandwich estimator³³. Imputed SNPs were analyzed using PLINK software³⁴ and genotype](#)
153 [imputation uncertainty was accounted for by using allelic dosage in PLINK.](#)
154 [Meta-analysis was performed using the fixed-effects inverse variance method based on the](#)
155 [regression \$\beta\$ estimates and standard errors from each study implemented in METAL³⁵. The](#)
156 [presence of heterogeneity between cohorts for the effect sizes of risk alleles was investigated](#)
157 [using the Cochran's \$Q\$ -test as implemented in METAL. To determine whether the genome-](#)
158 [wide significant signal at each locus with low LD in the same chromosomal region \(defined](#)
159 [as \$r^2 < 0.05\$ in a 750-kb region\) could be accounted for by a single SNP, we carried out](#)
160 [conditional analysis. Each cohort performed a genome-wide analysis for MODZT using](#)
161 [logistic regression adjusting for the top signal at each of the three associated regions to](#)
162 [determine whether potential second signals remained significant even after adjusting for these](#)
163 [variants. Results from each individual study were meta-analyzed to determine whether these](#)
164 [potential second signals were truly independent \(that is, if \$p < 5 \times 10^{-8}\$ \). In Iceland, being a](#)
165 [mother of twins was tested for association with the top alleles using logistic regression and](#)
166 [including age, age-squared and county of birth as covariates as described previously³⁶.](#)
167 [Associations between FSH level and genotype was assessed using linear regression as](#)
168 [described previously³⁶. To study the association of fertility measures and SNP genotype, we](#)
169 [used logistic regression \(having children\), Poisson log-linear regression \(number of children\),](#)
170 [or linear regression \(age at first child, age at last child, and average birth interval\). In these](#)
171 [analyses, birth cohort \(as a factor for each five year interval\), county of birth and six principal](#)
172 [components were included as covariates. Relatedness was controlled for by using genomic](#)

173 [control in all Icelandic association analyses](#)³⁷. [The combined p value of the meta-analysis and](#)
174 [the replication study was calculated using Fisher's combined probability test](#)³⁸.

176 **FSH Serum Level Measure**

177 [Serum levels of follicle-stimulating hormone \(FSH\) were measured in 2,411 men \(1,275](#)
178 [genotyped persons; 1,136 close relatives of genotyped individuals\) and 15,586 women \(9,738](#)
179 [genotyped persons; 5,848 close relatives of genotyped individuals\) referred to three clinics in](#)
180 [Iceland. FSH testing was undertaken primarily to investigate possible gonad impairment.](#)
181 [Hormone levels were measured by electrochemiluminescence immunoassay, using reagents](#)
182 [and analytical instruments from Roche Diagnostics GmbH, according to the manufacturer's](#)
183 [instructions.](#)

185 **In silico Functional Annotation**

186 [We used a number of publicly available bioinformatics tools and datasets to identify putative](#)
187 [functional effects of the top associated SNPs at each locus, including: Combined Annotation](#)
188 [Dependent Depletion \(CADD\)³⁹, HaploReg2 v4⁴⁰ and Variant effect predictor \(VEP\)⁴¹.](#)

190 **Gene Based Test**

191 [Results of the MODZT meta-analysis were used to perform a gene-based test of association](#)
192 [for the 63 candidate genes from Harris et al. using the *Knowledge-based mining system for*](#)
193 [Genome-wide Genetic studies \(KGG\) software Version 3.5^{42; 43}. This approach uses an](#)
194 [extended Simes test that integrates prior functional information and the meta-analysis](#)
195 [association results when combining the SNP p values within a gene in order to obtain an](#)
196 [overall association p value for each entire gene. As we tested for genetic association for 63](#)
197 [genes, the significance level was set at \$7.93 \times 10^{-4}\$ \(Bonferroni correction; 0.05/63\).](#)

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Polygenic Risk Scores (PRS)

Polygenic risk scores were calculated based on the results of the MODZT meta-analysis. Only markers having $r^2 > 0.9$ in all groups and $MAF > 0.01$ were included. To obtain effect sizes taking LD into account, the LDpred method developed by Vilhjálmsón and colleagues was used⁴⁴. As suggested by Vilhjálmsón et al, we calculated multiple sets of LD-modified effect sizes based on a grid of values for the fraction of causal markers ($\alpha = 0.0001, 0.0003, 0.001, 0.003, 0.01, 0.03, 0.1, 0.3, 1$). The resulting scores were then tested in a validation data set of Icelandic mothers of dizygotic twins. The score producing the most significant result in the validation data set were subsequently used to test for association with five fertility-related traits (“has children”, “number of children”, “age at first child”, “age at last child” and “average birth interval”).

Results

Genome-wide association results, replication and gene based analyses

The overall GWAS meta-analysis genomic control statistic (λ) was 1.01, indicating no appreciable inflation due to population structure. The quantile-quantile (Q-Q) plot of genome-wide p values showed a strong deviation from the null hypothesis of no association (Figure S1). The results are represented in the Manhattan plot (Figure 1). Three chromosomal regions contained genome-wide significant SNPs ($p < 5 \times 10^{-8}$). Twenty-two SNPs on chromosome 11p13 showed genome-wide significant associations. Of these, the strongest signal (rs11031006, $p = 1.54 \times 10^{-9}$ and OR 1.41, 95% CI 1.29-1.53) lies in the region 5' of the transcription start site of *FSHB* (encoding the FSH, beta polypeptide, [MIM: 136530]). A second locus was represented by five SNPs on *Iq42.13* covering an intergenic region flanked by *PGBD5* (MIM: 616791) and *COG2* (MIM: 606974), ($p = 1.23 \times 10^{-8}$, OR 1.43, 95% CI

223 1.31-1.55 for strongest signal, rs12064669). A third locus on *15q22.33*, mapped to the first
224 intron of *SMAD3* (SMAD family member 3, [MIM: 603109]) (rs17293443, $p = 1.57 \times 10^{-8}$
225 and OR 1.27, 95% CI 1.19-1.35). No significant heterogeneity in SNP effects was observed
226 across cohorts for the top SNPs ($p > 0.1$, Cochran's Q test); Tables [43](#) and [S23](#)). The
227 regional association plots for these loci are shown in figure S2. After conditioning on the top
228 SNPs at each locus, no secondary signals were observed (all $p > 0.05$; Tables [S43](#), [S54](#) and
229 [S65](#)). We sought validation of these three top signals in an independent replication study from
230 Iceland (deCODE) totaling 3,597 mothers of twins and 297,348 controls. The *FSHB*
231 (rs11031006, $p = 3 \times 10^{-3}$, OR 1.14, 95% CI 1.06-1.22) and *SMAD3* (rs17293443, $p = 1.44 \times$
232 10^{-4} , OR 1.15, 95% CI 1.07-1.23) loci replicated, but rs12064669 ($p=0.88$) was not confirmed
233 (Table [43](#) and Figure [S32](#)). We also investigated whether any of the proposed 63 candidate
234 genes¹⁴ was associated with human DZ twinning. In a gene-based test, five genes
235 demonstrated a nominally significant association ($p < 0.05$; *BMPRIA* [MIM: 601299],
236 *BMPR1B* [MIM: 603248], *IGF1* [MIM: 147440], *FSHB* and *FSHR* [MIM: 136435]). After
237 correction for multiple testing, only *FSHB* remained significant (Table [S76](#)).

Functional in silico annotation of associated variants

240 We explored plausible functional effects of our associated variants using Combined
241 Annotation Dependent Depletion (CADD).³⁹ The *FSHB* SNP rs11031006 had a high Phred
242 scaled C-score (22.4), indicating that it is among the top 1% of SNPs in the human genome
243 most likely to have a functional effect. The Phred score for the *SMAD3* SNP rs17293443 was
244 only 2.71 indicating that it is among the bottom 50% of SNPs in the human genome likely to
245 have a functional effect (Table [S407](#)). Examination of individual constituents of the CADD
246 scores showed particularly high conservation-based scores for rs11031006 (Figure [S63](#)). To
247 further investigate possible functional effects we examined data from the ENCODE project.⁴⁵

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248 The *FSHB* SNP rs11031006 alters the sequence of 11 protein-binding motifs including that of
249 the Estrogen receptor alpha, indicating a possible effect on hormonal feedback inhibition
250 (Table S448). Variant Effect Predictor (VEP)⁴¹ identified rs17293443 as a regulatory region
251 variant within a promoter flanking region (ENSR00000410126) (Figure S74). SNP
252 rs17293443 was contained in a *DNase I* hypersensitive site suggesting open chromatin,
253 although it did not alter any of the transcription factor binding sites present in the promoter
254 flanking region. Together, these data indicate that rs11031006 and rs17293443 may have
255 direct functional roles.

256 **SNPs associated with FSH serum level**

257 We analyzed serum FSH measurements from 17,997 genotyped Icelanders and their close
258 relatives. SNP rs11031006 was significantly associated with higher serum FSH levels ($p = 2.3$
259 $\times 10^{-10}$), with each G allele conferring an increase in FSH level of about 0.11 SD units (Figure
260 S54). Notably, the allele (rs11031006-G) conferring the strongest association with FSH levels
261 in the FSH GWAS is the same allele that conferred the greatest chance of having DZ twins in
262 our MODZT GWAS. No association was seen between the *SMAD3* signal and serum FSH
263 levels ($p = 0.30$).

265 **Relative risk of twin birth and polygenic risk score prediction of DZ twinning and** 266 **fertility measures**

267 We estimated that the relative risk of a twin birth, based on approximately 90,000 births in
268 Iceland between 1950 and 1991, was increased by 18% for each maternal rs11031006-G
269 allele and by 9% for each rs17293443-C allele (Table S8). A higher polygenic risk score
270 (PRS) for DZ twinning, calculated based on the results of the DZ twinning GWAS, was
271 significantly associated with DZ twinning in our independent Icelandic cohort ($p = 0.001$)
272 (Table S7). A higher PRS was also associated with a higher likelihood of having children ($p =$

273 0.01), higher lifetime number of children ($p = 0.03$), and an earlier age at first child ($p = 0.02$)
274 (Table S98). A re-calculated PRS, excluding the 1 Mb regions surrounding the two replicated
275 variants, remained associated with DZ twinning ($p = 0.02$) and with the likelihood of having
276 children ($p = 0.03$). These results reflect the polygenic contribution to the susceptibility to DZ
277 twinning and its association with greater reproductive ability.

279 SNPs associated with female reproduction traits

280 Table 24 reports on the two loci robustly implicated in DZ twinning and other reproductive
281 traits in women. Consistent with its effects on higher circulating FSH levels, the rs11031006-
282 G allele is also associated with earlier age at menarche,^{22; 46} earlier age at first child and
283 higher total lifetime number of children, lower risk of polycystic ovary syndrome (PCOS),²³
284 and earlier age at natural menopause.^{24; 47} Also, the DZ twinning SNP rs11031006 is
285 correlated with a variant (rs10835638, *FSHB*-211G>T) located in the promoter of *FSHB* ($r^2 =$
286 0.62) that is associated with timing of breast development in girls.⁴⁸ In contrast, the
287 rs17293443-C allele in *SMAD3* was associated only with a later age at last child (Figure S65).

288 ~~Functional in silico annotation of associated variants~~

289 ~~We explored plausible functional effects of our associated variants using Combined~~
290 ~~Annotation Dependent Depletion (CADD).³⁸ The *FSHB* SNP rs11031006 had a high Phred~~
291 ~~scaled C score (22.4), indicating that it is among the top 1% of SNPs in the human genome~~
292 ~~most likely to have a functional effect. The Phred score for the *SMAD3* SNP rs17293443 was~~
293 ~~only 2.71 indicating that it is among the bottom 50% of SNPs in the human genome likely to~~
294 ~~have a functional effect (Table S10). Examination of individual constituents of the CADD~~
295 ~~scores showed particularly high conservation based scores for rs11031006 (Figure S6). To~~
296 ~~further investigate possible functional effects we examined data from the ENCODE project.⁴⁴~~
297 ~~The *FSHB* SNP rs11031006 alters the sequence of H1 protein binding motifs including that of~~

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298 ~~the Estrogen receptor alpha, indicating a possible effect on hormonal feedback inhibition~~
299 ~~(Table S11). Variant Effect Predictor (VEP)⁴⁰ identified rs17293443 as a regulatory region~~
300 ~~variant within a promoter flanking region (ENSR00000410126) (Figure S7). SNP rs17293443~~
301 ~~was contained in a DNase I hypersensitive site suggesting open chromatin, although it did not~~
302 ~~alter any of the transcription factor binding sites present in the promoter flanking region.~~
303 ~~Together, these data indicate that rs11031006 and rs17293443 may have direct functional~~
304 ~~roles.~~

305 Discussion

306 Here we report the [first](#) compelling evidence that sequence variation at the *FSHB* and
307 *SMAD3* loci increases the odds of DZ twinning. A number of studies in mothers of DZ twins,
308 but not all,⁴⁹ have found higher FSH levels responsible for multiple follicle growth.⁵⁰ The
309 associations of *FSHB* rs11031006-G with earlier ages at breast development, menarche,
310 menopause, first child, and higher lifetime parity indicates that this locus plays an important
311 role in multiple reproductive aspects. Female carriers of rs11031006-G likely may have a
312 more advanced depletion of the ovarian follicular pool and hence would have an increased
313 risk of premature ovarian failure (POF, [MIM: 612964]). Indeed advanced ovarian aging is a
314 recognized feature of familial DZ twinning, with reported lower levels of anti-Mullerian
315 Hormone (AMH) a marker of lower ovarian primordial follicular reserve.⁵¹ The rs17293443-
316 C allele in *SMAD3* also increases chances of DZ twinning, but this effect appears independent
317 of circulating FSH levels.

318 Of the 63 suggested candidate genes for twinning, only *FSHB* was associated in gene-
319 based tests [after correcting for multiple testing](#). Emerging data in human and non-human
320 primates suggest that mechanisms underlying multiple ovulation may differ among species,
321 explaining why in some species twinning is accompanied by unique evolutionary adaptations
322 enabling offspring's survival, such as the marmoset diminutive fetal size in a simplex uterus,

323 and subsequent alloparenting.¹⁴ Conversely, in humans, multiple gestations remain an
324 independent risk factor for preterm birth, pregnancy loss, and fetal growth restriction. Thus
325 loci common to species and strains with higher rates of DZ twinning will not necessarily be
326 shared, as the rate, complications, and future reproductive fitness of those twin gestations
327 differ.

328 The third genome-wide hit from the discovery was not replicated in Iceland. SNP
329 rs120644669 is located 149kb from the angiotensinogen (*AGT*, [MIM: 106150]), which
330 influences ovulatory capacity in mice,⁵² and 89kb from component of oligomeric golgi
331 complex 2 (*COG2*) affecting protein glycosilation, which regulates the biological activity of
332 the pituitary gonadotrophins.⁵³

333 Genetic variants near *FSHB* (rs11031005 and rs11031002, which are highly correlated
334 with rs11031006) are reportedly associated not only with higher serum FSH, but also lower
335 LH levels.⁵⁴ This agrees with the known response of the gonadotrophic cell to a high
336 frequency hypothalamic pulsatile GnRH signal that reciprocally controls secretion of both
337 hormones.⁵⁵ Furthermore, under normal conditions suppression of FSH in the early follicular
338 phase and higher LH levels in late phase typically favor mono-ovulation in the human.⁵⁶ In a
339 study aiming to identify genetic predictors for IVF success or IVF-controlled ovarian
340 stimulation (COS), rs611246 located in *FSHB* ($r^2 = 0.3$ with rs11031006), was reported
341 significantly associated with measured early follicular phase FSH values and also with the
342 probability of clinical pregnancy, suggesting that these genetic variants are potential predictor
343 candidates that could be considered in clinical ovarian reserve and function assessment in
344 assisted reproduction.⁵⁷

345 A recent linkage study in cattle reported only one strong signal ($p < 1 \times 10^{-28}$) for
346 ovulation rate in a region spanning *SMAD3*, *SMAD6* (MIM: 602931) and *IQCH* (MIM:
347 612523).⁵⁸ *SMAD3* encodes one of a family of proteins that function as signal transducers and

348 transcriptional modulators that mediate multiple signaling pathways. Observations in mice
349 have established an essential role for Smad3 in mediating TGFbeta and activin signals in the
350 ovarian granulosa cell and also in the pituitary⁵⁹ to maintain a favorable environment for
351 oocyte maturation.⁶⁰ Smad3 is strongly expressed in the human ovary, where it promotes
352 granulosa cell proliferation and steroidogenesis possibly by upregulating gonadotrophin
353 receptor signaling pathways.⁶¹ Thus, sequence variation in *SMAD3* may increase chance of
354 DZ twinning by increasing responsiveness to FSH. Understanding the role of *SMAD3* will
355 offer novel opportunities to optimize responsiveness and minimize risk among assisted
356 reproduction technology (ART) recipients for example through adjustment of hormonal
357 stimulation and thus contributes to prevention of life-threatening ovarian hyper stimulation
358 syndrome in hyper responding female carriers of the rs17293443-C allele and conversely
359 prevention of a poor response in [patients-individuals](#) with allelic variants that lead to a poor
360 response.

361 The potential applications of this work in reproductive medicine are multifold. Firstly,
362 it reveals a well-defined set of loci for DZ twinning in the general population. Secondly,
363 twinning is associated with common perinatal morbidities such as preterm birth, discordant
364 twin growth, latter prenatal asymmetric intrauterine growth restriction, and placental
365 abruption. Multiple gestations are also related to a higher prevalence of maternal morbidities
366 such as preeclampsia, postpartum hemorrhage, and ensuing complications. By understanding
367 the genetic basis of DZ twinning, we concomitantly identify loci conferring susceptibility (or
368 conversely resistance) to these prevalent perinatal comorbidities. The involvement of *SMAD3*
369 is of relevance in the light of a possible phenotype in male DZ twins namely the repeatedly
370 reported more frequent occurrence of testicular seminoma, a gonadal tumor in which the
371 mitogenic cyclin D2 is overexpressed.^{62; 63} In animal experimental studies knockout of
372 *SMAD3* significantly attenuates cyclin D2 and tumor proliferation.⁶⁴ This indicates to the

373 likely clinical potential of *SMAD3* stretching beyond that of fertility physiology and
374 treatment.

375 [This study](#) ~~The sequence variants found to influence spontaneous DZ twinning and~~
376 ~~their relationship with other fertility related measures,~~ provides important insights into
377 ovarian functioning and the control of natural multiple follicle growth and reproductive aging.
378 This has important implications for fertility, including improved outcome prediction and
379 novel avenues of fertility treatment. [Other strengths include the rigorous phenotype inclusion](#)
380 [of mothers of DZ twins with a documented history of spontaneous DZ twinning. It is worth](#)
381 [mentioning that analyses without excluding mothers who conceived their twins with hormone](#)
382 [induction of multiple ovulation or other ART, did not yield to any genome-wide significant](#)
383 [results \(results not shown\). We thus recommend twins registries collect data on mode of](#)
384 [pregnancy of twins. There are also some limitations. The current effort focused on unraveling](#)
385 [the genetic basis of DZ twinning in European-ancestry populations only. However, the](#)
386 [highest incidence of DZ twins was reported in the Nigerian population and some other](#)
387 [countries from Africa. Next steps in unraveling the genetic cause of DZ twinning need to](#)
388 [include mothers of DZ twins originating from these regions, and if possible also from regions](#)
389 [where DZ twinning is a rare trait, such as Japan².](#)

390 In summary, we identified *FSHB* and the ~~novel~~-*SMAD3* locus as maternal
391 susceptibility loci for DZ twinning. These loci are also significantly and specifically
392 associated with several other aspects of reproductive capacity and health.

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Supplemental Data

Supplemental data include ~~seven~~six figures and ~~four~~eight tables

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Web Resources

1000GenomesProject, <ftp://ftp-trace.ncbi.nih.gov/1000genomes/ftp/release/20110521/>

CADD, <http://cadd.gs.washington.edu/>

HaploReg, <http://www.broadinstitute.org/mammals/haploreg/haploreg.php>

Variant Effect Predictor (VEP), <http://www.ensembl.org/info/docs/tools/vep/index.html>

Knowledge-based mining system for Genome-wide Genetic studies (KGG),

<http://grass.cgs.hku.hk/limx/kgg/>

ENCODE, <https://genome.ucsc.edu/ENCODE/>

OMIM, <http://www.omim.org/>

LDpred, https://bitbucket.org/bjarni_vilhjalmsson/ldpred

MACH, <http://www.sph.umich.edu/csg/abecasis/MACH/>

Minimac, <http://genome.sph.umich.edu/wiki/Minimac>

Beagle, <https://faculty.washington.edu/browning/beagle/b3.html>

PLINK, <http://pngu.mgh.harvard.edu/~purcell/plink/>

LocusZoom, <http://csg.sph.umich.edu/locuszoom/>

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Figures titles and legends

Figure 1. Manhattan Plot for the genome-wide association meta-analysis of mothers of dizygotic twins. SNPs are plotted on the *x*-axis according to their position on each chromosome against the significance of the association on the *y*-axis (shown as $-\log_{10} p$

values). The dotted red line denotes $p=5\times 10^{-8}$ statistical significance. The arrows point to the chromosomal region that reached genome-wide significance level.

[Figure 2. Forest plots depicting risk allele odds ratio \(OR\) estimates at \(A\) rs11031006 \(near FSHB\), \(B\) rs17293443 \(SMAD3\) and \(C\) rs12064669 from each study, the meta-analysis, deCODE \(replication\) and the overall combined results. Black squares indicate the OR and horizontal lines represent the 95% CIs. The combined results are indicated by the black diamond.](#)

Tables titles and legends

[Table 1. Number of cases and controls included in the meta-analysis and replication; mean maternal age at delivery](#)

[Table 2. Sample characteristics for fertility measures traits](#)

Table [43](#). Genome-wide significant loci in the meta-analysis of mothers of spontaneous DZ twins ($n = 1,980$) and replication ($n = 3,597$) versus screened controls

Table [24](#). Association results for the implicated DZ twinning SNPs in fertility-related measures

Table 1. Number of cases and controls included in the meta-analysis and replication; mean maternal age at delivery

<u>Cohort</u>	<u>No. of subjects</u>		<u>Mean maternal age at delivery (SD)</u>
	<u>Cases</u>	<u>Controls</u>	
<u>Netherlands</u>	<u>806</u>	<u>4,535</u>	<u>29.7 (4.1)</u>
<u>Australia</u>	<u>606</u>	<u>6,556</u>	<u>29.8 (4.8)</u>
<u>Minnesota</u>	<u>568</u>	<u>1,862</u>	<u>28.3 (4.5)</u>
<u>Total discovery</u>	<u>1,980</u>	<u>12,953</u>	<u>29.3 (4.5)</u>
<u>Replication</u>			
<u>Iceland</u>	<u>3,597 (1,356 genotyped; 2,241 <i>in silico</i>)</u>	<u>297,348 (76,342 genotyped; 221,006 <i>in silico</i>)</u>	<u>30.3 (5.8)</u>

Table 2. Sample characteristics for measure of fertility

<u>Traits</u>	<u>Sample size</u>	<u>Mean (SD)</u>
<u>Age at menarche</u>	<u>259,247</u>	<u>±</u>
<u>Age at menopause</u>	<u>69,360</u>	<u>±</u>
<u>FSH levels</u>	<u>17,997</u>	<u>=</u>
<u>Has children</u>	<u>41,946</u>	<u>0.94²</u>
<u>Number of children</u>	<u>41,946</u>	<u>2.9 (1.6)</u>
<u>Age at first child</u>	<u>41,946</u>	<u>23.0 (4.7)</u>
<u>Age at last child</u>	<u>41,946</u>	<u>32.1 (5.5)</u>
<u>Average birth interval</u>	<u>41,946</u>	<u>4.7 (2.5)</u>
<u>Polycystic ovary syndrome</u>	<u>5,184 cases / 82,759 controls</u>	<u>±</u>

¹ this is a large meta-analysis of summary statistics and sample mean is not available

² fraction with children;

Table 13. Genome-wide significant loci in the meta-analysis of mothers of spontaneous DZ twins (n=1,980) and replication (n=3,597) versus screened controls

SNP	Locus	Position ^a	Gene	Annotation	Risk allele	Meta-analysis			Replication			Combined
						RAF	OR (95%CI)	p	RAF	OR (95% CI)	p	p
rs11031006	<i>11p14.1</i>	30226528	<i>FSHB</i>	5' upstream	G	0.85	1.41 (1.29-1.53)	1.54×10 ⁻⁹	0.85	1.14 (1.06-1.22)	3×10 ⁻³	1.25×10 ⁻¹⁰
rs17293443	<i>15q22.33</i>	67437863	<i>SMAD3</i>	Intron	C	0.24	1.27 (1.19-1.35)	1.57×10 ⁻⁸	0.21	1.15 (1.07-1.23)	1.44×10 ⁻⁴	6.29×10 ⁻¹¹
rs12064669	<i>1q42.13</i>	230688643		Intergenic	C	0.10	1.43 (1.31-1.55)	1.23×10 ⁻⁸	0.07	1.01 (0.89-1.13)	0.88	2.09×10 ⁻⁷

^aSNP position according to NCBI Human Genome Build 37; RAF, risk allele frequency; OR, odd ratio; 95% CI, 95% confidence interval

Table 24. Association results for the implicated DZ twinning SNPs in fertility-related measures

DZ twinning and fertility-related measure	rs11031006-G allele (near <i>FSHB</i>)			rs17293443-C allele (in <i>SMAD3</i>)		
	Effect	OR (95% CI)/ <u>Beta (95% CI)</u>	p value	Effect	OR (95% CI)/ <u>Beta (95% CI)</u>	p value
DZ twinning ^a	Increase	<u>1.41 (1.29, 1.53)</u>	1.54×10 ⁻⁹	Increase	<u>1.27 (1.19, 1.35)</u>	1.57×10 ⁻⁸
FSH levels (<u>SD units</u>) ^b	Increase	<u>0.11 (0.078, 0.15)</u>	2.3×10 ⁻¹⁰	–	<u>0.016 (-0.014, 0.045)</u>	0.30
Age at menarche (<u>years</u>) ^{c,d}	Earlier	<u>-0.04 (-0.012, 0.011)</u>	8.5×10 ⁻¹⁰	Earlier	<u>-0.001 (-0.127, 0.010)</u>	0.84
Age at menopause (<u>years</u>) ^e	Earlier	<u>-0.2165 (-0.065, 0.052)</u>	8.5×10 ⁻¹⁴	–	<u>0.009 (-0.048, 0.049)</u>	-0.71
Has children (yes/no) ^f	–	<u>1.07 (0.98, 1.17)</u>	0.12	–	<u>0.96 (0.89, 1.04)</u>	0.34
Number of children ^f	Increase	<u>0.014 (0.00091, 0.27)^h</u>	0.03	–	<u>0.0048 (-0.0062, 0.016)^h</u>	0.39
Age at first child (<u>years</u>) ^f	Earlier	<u>-0.20 (-0.31, -0.086)</u>	5.3×10 ⁻⁴	–	<u>0.032 (-0.065, 0.13)</u>	0.51
Age at last child (<u>years</u>) ^f	Earlier	<u>-0.097 (-0.21, 0.015)</u>	0.08	Later	<u>0.14 (0.043, 0.24)</u>	4.7×10 ⁻³
Average birth interval (<u>years</u>) ^f	–	<u>0.015 (-0.037, 0.057)</u>	0.57	–	<u>0.0051 (-0.040, 0.041)</u>	0.82

		0.067			0.050	
PCOS ^g	Decrease	0.90 (0.84, 0.95)	$+2.39 \times 10^{-49}$	-	1.00 (0.96, 1.05)	0.78

^aMODZT GWAS; ^bdeCODE sample of 17,997 individuals with FSH levels; ^{c,d}[Lunetta et al. and Perry et al. \(ref 21 and 221\) and Perry et al. \(ref 22\)](#); ^eDay et al. (ref [2424](#)) and [Stolk et al. \(ref 25\)](#); ^fdeCODE sample of 41,946 women; ^gPolycystic ovary syndrome, [Hayes-Day et al. \(ref 223\)](#); ^h[factor of increase from log-linear model.](#)