

**Body-mass index and all-cause mortality:****Individual-participant-data meta-analysis of 239 prospective studies in four continents**The Global BMI Mortality Collaboration (investigators listed in **eAppendix 1**)

Emanuele Di Angelantonio<sup>\*1</sup>, MD; Shilpa N Bhupathiraju<sup>\*2</sup>, PhD; David Wormser<sup>\*1</sup>, PhD; Pei Gao<sup>\*1,3</sup>, PhD; Stephen Kaptoge<sup>\*1</sup>, PhD; Amy Berrington de Gonzalez<sup>\*4</sup>, PhD; Benjamin J. Cairns<sup>\*5</sup>, PhD; Prof Rachel Huxley<sup>\*6</sup>, PhD; Chandra L. Jackson<sup>\*7</sup>, PhD; Grace Joshy<sup>\*8</sup>, PhD; Sarah Lewington<sup>\*5</sup>, DPhil; Prof JoAnn E Manson<sup>\*2,7</sup>, MD; Neil Murphy<sup>\*9</sup>, PhD; Alpa V. Patel<sup>\*10</sup>, PhD; Prof Jonathan M. Samet<sup>\*11</sup>, MD; Prof Mark Woodward<sup>\*5,12,13</sup>, PhD; Wei Zheng<sup>\*14</sup>, MD; Maigen Zhou<sup>\*15</sup>, MSc; Narinder Bansal<sup>1</sup>, PhD; Aurelio Barricarte<sup>16,17</sup>, MD; Brian Carter<sup>10</sup>, MPH; Prof James R Cerhan<sup>18</sup>, MD; Prof Rory Collins<sup>5</sup>, FRS; Prof George Davey Smith<sup>19</sup>, MD; Xianghua Fang<sup>20</sup>, PhD; Prof Oscar H. Franco<sup>21</sup>, MD; Prof Jane Green<sup>5</sup>, DPhil; Jim Halsey<sup>5</sup>, BSc; Janet S. Hildebrand<sup>10</sup>, MPH; Keum Ji Jung<sup>22</sup>, MPH; Rosemary J. Korda<sup>8</sup>, PhD; Dale F. McLerran<sup>23</sup>, MS; Steven C. Moore<sup>4</sup>, PhD, MPH; Linda M. O'Keefe<sup>1</sup>, PhD; Ellie Paige<sup>1</sup>, PhD; Anna Ramond<sup>1</sup>, DPharm; Prof Gillian K. Reeves<sup>5</sup>, PhD; Betsy Rolland<sup>4</sup>, PhD, MPH; Carlotta Sacerdote<sup>24</sup>, PhD; Prof Naveed Sattar<sup>25</sup>, FRCP; Eleni Sofianopoulou<sup>1</sup>, PhD; Prof June Stevens<sup>26</sup>, PhD; Michael Thun<sup>10</sup>, MD; Prof Hirotsugu Ueshima<sup>27</sup>, MD; Ling Yang<sup>5</sup>, PhD; Young Duk Yun<sup>28</sup>, MD; Peter Willeit<sup>1,29</sup>, PhD; Prof Emily Banks<sup>\*8</sup>, PhD; Prof Valerie Beral<sup>\*5</sup>, FRS; Prof Zhengming Chen<sup>\*5</sup>, MD; Susan M. Gapstur<sup>\*10</sup>, PhD; Marc J. Gunter<sup>\*30</sup>, PhD; Patricia Hartge<sup>\*4</sup>, ScD; Prof Sun Ha Jee<sup>\*23</sup>, PhD; Prof Tai-Hing Lam<sup>\*31</sup>, FRCP; Prof Richard Peto<sup>\*5</sup>, FRS; Prof John D. Potter<sup>\*32</sup>, PhD; Prof Walter C. Willett<sup>\*2,7</sup>, MD; Prof Simon G. Thompson<sup>\*1</sup>, FMedSci; Prof John Danesh<sup>\*1</sup>, FMedSci; Prof Frank B. Hu<sup>\*2,7</sup> MD.

\* Equal contribution

<sup>1</sup> University of Cambridge, Cambridge, UK<sup>2</sup> Harvard T.H. Chan School of Public Health, Boston, Massachusetts, USA<sup>3</sup> Peking University, Beijing, China<sup>4</sup> National Cancer Institute, Bethesda, Maryland, USA<sup>5</sup> University of Oxford, Oxford, UK<sup>6</sup> University of Queensland, Brisbane, Australia<sup>7</sup> Harvard Medical School, Boston, Massachusetts, USA<sup>8</sup> Australian National University, Canberra, Australia<sup>9</sup> Imperial College London, London, UK<sup>10</sup> American Cancer Society, Atlanta, Georgia, USA<sup>11</sup> University of Southern California, Los Angeles, California, USA<sup>12</sup> The George Institute for Global Health, University of Sydney, Australia<sup>13</sup> Johns Hopkins University, Baltimore, Maryland, USA<sup>14</sup> Vanderbilt University Medical Center, Nashville, Tennessee, USA<sup>15</sup> Chinese Center for Disease Control and Prevention, Beijing, China<sup>16</sup> Navarre Public Health Institute, Pamplona, Spain<sup>17</sup> Consortium for Biomedical Research in Epidemiology and Public Health, Spain<sup>18</sup> Mayo Clinic, Rochester, Minnesota, USA<sup>19</sup> MRC Integrative Epidemiology Unit, University of Bristol, Bristol, UK<sup>20</sup> Capital Medical University, Beijing, China<sup>21</sup> Erasmus MC, University Medical Center Rotterdam<sup>22</sup> Yonsei University, Seoul, Korea<sup>23</sup> Fred Hutchinson Cancer Research Center, Seattle, USA<sup>24</sup> University of Turin, Center for Cancer Prevention, Turin, Italy<sup>25</sup> University of Glasgow, Glasgow, UK<sup>26</sup> University of North Carolina, Chapel Hill, North Carolina, USA<sup>27</sup> Shiga University of Medical Science, Shiga, Japan<sup>28</sup> Health Insurance Policy Research Institute, Seoul, Korea<sup>29</sup> Medical University Innsbruck, Innsbruck, Austria<sup>30</sup> International Agency for Research on Cancer, Lyon, France<sup>31</sup> School of Public Health University of Hong Kong, Hong Kong, China<sup>32</sup> Massey University, Wellington, New Zealand

57  
58 Correspondence: Professor John Danesh  
59 Department of Public Health and Primary Care  
60 University of Cambridge  
61 Cambridge, England  
62 gbmc@phpc.cam.ac.uk

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67 Abstract 298 words

68 Text 3307 words

69 2 tables and 4 figures

70 39 eTables/eFigures/eAppendices

71 **SUMMARY**

72 **Background:** Overweight and obesity are increasing worldwide. To help assess their relevance to  
73 mortality in different populations we conducted individual-participant-data meta-analyses of the  
74 prospective studies of body mass index (BMI), limiting confounding and reverse causality by  
75 restricting analyses to never-smokers and excluding prior disease and the first 5 years of follow-  
76 up.

77 **Methods:** Of 10,625,411 participants in Asia, Australia/New Zealand, Europe, and North America  
78 from 239 prospective studies (median follow-up 13.7 [IQR: 11.4-14.7] years), 3,951,455 in 189  
79 studies were never-smokers without specific chronic diseases at recruitment who survived 5 years,  
80 of whom 385,879 died. The primary analyses are of these deaths, using age and sex-adjusted  
81 hazard ratios (HRs).

82 **Findings:** All-cause mortality was minimal (HR=1) at BMI (kg/m<sup>2</sup>) 20-25, and increased  
83 significantly both just below this range (BMI 18.5-<20: HR=1.13, 95%CI 1.09-1.17; BMI 15-  
84 <18.5: HR=1.51, 1.43-1.59) and throughout the overweight range just above it (BMI 25-<27.5:  
85 HR=1.07, 1.07-1.08; BMI 27.5-<30: HR=1.20, 1.18-1.22). Continuing upwards, HRs for obesity  
86 grade I, II, and III (BMI 30-<35, 35-<40, 40-<60) were 1.45 (1.41-1.48), 1.94 (1.87-2.01), and  
87 2.76 (2.60-2.92), respectively. For BMI>25, mortality increased approximately log-linearly with  
88 BMI; HR per 5 units higher BMI was 1.31 (1.29-1.33) in all regions; 1.39 (1.35-1.43) in Europe;  
89 1.29 (1.26-1.32) in North America, 1.39 (1.34-1.44) in East Asia, and 1.31 (1.27, 1.35) in  
90 Australia/New Zealand. This HR per 5 units higher BMI (for BMI>25) was greater in younger than  
91 older people (1.52 at 35-49 years vs 1.21 at 70-89 years;  $P_{\text{heterogeneity}} < 0.0001$ ), greater in men  
92 than women (1.51 vs 1.30;  $P_{\text{heterogeneity}} < 0.0001$ ), but similar in studies with self-reported and  
93 measured BMI.

94 **Interpretation:** The associations of both overweight and obesity with higher all-cause mortality  
95 were broadly consistent in four continents. This supports strategies to combat the entire spectrum  
96 of excess adiposity in many populations.

97 **Funding:** UK MRC, BHF, NIHR; US NIH

98 **INTRODUCTION**

99  
100 The worldwide prevalence of overweight and obesity is high and increasing.<sup>1</sup> The World Health  
101 Organization (WHO) estimates that more than 1·3 billion adults worldwide are overweight (defined  
102 by WHO as a body mass index [BMI weight in kg/the square of height in m] of 25-<30 kg/m<sup>2</sup>), and  
103 a further 600 million are obese (BMI ≥30 kg/m<sup>2</sup>).<sup>2</sup> Analyses of large-scale prospective studies with  
104 prolonged follow-up generally indicate that both overweight and obesity are associated with  
105 increased mortality, as is underweight (defined conservatively by the WHO as BMI <18.5 kg/m<sup>2</sup>).  
106 However, it is not known how such associations vary across major global regions, an uncertainty  
107 relevant to international strategies on overweight and obesity.<sup>3</sup> Most previous analyses have  
108 focused on people living in one particular country or continent,<sup>4-11</sup> even though relationships of  
109 overweight and underweight might differ from one population to another.

110  
111 Estimation of the relationships between BMI and mortality in various populations can help assess  
112 the adverse physiological effects of excessive adiposity (and the adverse physiological effects of  
113 various determinants of low BMI). However, reliable estimates of the relevance of BMI to mortality  
114 need to limit the effects of reverse causality, because chronic disease and smoking can themselves  
115 affect BMI. To help achieve more valid estimates, prospective studies of BMI and mortality should  
116 where possible exclude smokers, exclude participants who already had some chronic disease at  
117 recruitment that could affect BMI, and those dying within 5 years of recruitment.<sup>12-15</sup>

118  
119 The Global BMI Mortality Collaboration was established to provide a standardised comparison of  
120 associations of BMI with mortality across different populations. It includes individual-participant  
121 data on 10·6 million adults in 239 prospective cohort studies in 32 countries, mainly located in  
122 Australia/New Zealand (NZ), East Asia, Europe, or North America, about 4 million of whom were  
123 non-smokers without chronic disease at recruitment who were still being followed up 5 years  
124 afterwards.

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## METHODS

In 2013, over 500 investigators (**eAppendix 1**) from over 300 institutions in 32 countries agreed an analysis plan for combining individual-participant data from contributing studies. This pre-specified analysis plan is provided in **eAppendix 2**. The goal was to produce reliable estimates of associations of overweight and obesity with all-cause mortality using data from studies in several regions. The pre-specified analysis methods were designed to maximize the internal validity by reducing the scope for bias. This paper follows PRISMA-IPD reporting guidelines (**eAppendix 3**).<sup>16</sup>

### Data Sources

We sought data from large prospective studies ( $\geq 100,000$  participants at baseline) or large multi-cohort consortia (total  $\geq 100,000$  participants at baseline). We identified studies from 1970 to January 2015 through systematic literature searches and discussion with investigators (**eAppendix 4**). Prospective cohort studies or consortia thereof were eligible if they: 1) had information on weight, height, age and sex; 2) did not select participants on the basis of having previous chronic disease; 3) recorded overall or cause-specific deaths; and 4) had accrued  $\geq 5$  years of median follow-up. We identified only two eligible studies that were unable to contribute (**eFigure 1**).<sup>17,18</sup> **eTables 1-2** provide details of studies. The contributing studies classified deaths according to the primary cause (or, in its absence, the underlying cause), on the basis of coding from the *International Classification of Diseases*, revisions 8 through 10, to at least three digits (**eTable 3**), or according to study-specific classification systems. Ascertainment of outcomes was based on death certificates, supplemented in some studies by additional data.

### Statistical methods

**eFigure 1** describes inclusion and exclusion criteria. We excluded participants with a BMI  $< 15$  or  $\geq 60 \text{ kg/m}^2$ , or baseline age  $< 20$  or  $\geq 90$  years. To limit residual confounding by smoking and bias due to effects of pre-existing disease on baseline BMI (i.e., reverse causality), the primary analysis was restricted to never-smokers without certain known chronic diseases at baseline (e.g., cardiovascular disease, cancer, or respiratory diseases), and omitted the first 5 years of follow-up.

Associations of all-cause mortality with BMI depend not only on the associations of specific causes of death with BMI in different regions (which might differ quantitatively), but also on how relatively

157 common each specific cause of death is in the particular region (which can differ substantially by  
158 region and over time). Hence, the association of all-cause mortality with BMI may differ in regions  
159 with different underlying mortality patterns. Therefore, the pre-specified primary analysis was  
160 stratified by 5 major geographical regions, 3 with extensive data (East Asia, Europe, North  
161 America) and 2 with more limited data (Australia/NZ and South Asia). Data from some or all  
162 regions are shown separately, in the main text or the extensive online Supplementary Analyses.

163  
164 Each study (or consortium of studies) analysed individual-participant data according to a common  
165 analytical plan (SASv9.3 [SAS Institute, Cary, NC] or StataV12 [StataCorp, College Station, TX])  
166 provided by the coordinating centres. These separate results were then meta-analysed at  
167 Cambridge University, UK. To facilitate standardised comparisons with other meta-analyses, we  
168 calculated hazard ratios (HRs) for mortality in the 6 WHO-defined baseline BMI categories:  
169 underweight ( $15.0 < 18.5 \text{ kg/m}^2$ ), normal ( $18.5 < 25 \text{ kg/m}^2$ , the reference category for analyses of  
170 6 BMI groups), overweight ( $25 < 30 \text{ kg/m}^2$ ), and obesity grades I ( $30 < 35 \text{ kg/m}^2$ ), II ( $35 < 40$   
171  $\text{kg/m}^2$ ), and III ( $40 < 60 \text{ kg/m}^2$ ).<sup>19</sup> As, however, most people are of normal weight or overweight,  
172 these two categories were subdivided, yielding 9 fine groups:  $15.0 < 18.5$ ,  $18.5 < 20$ ,  $20 < 22.5$ ,  
173  $22.5 < 25.0$  (the reference category for analyses of 9 BMI groups),  $25.0 < 27.5$ ,  $27.5 < 30$ ,  $30 <$   
174  $35.0$ ,  $35.0 < 40$ , and  $40 < 60 \text{ kg/m}^2$ .

175  
176 Study-specific log HRs in particular BMI categories were pooled by inverse-variance-weighted  
177 random-effects meta-analyses (an extension of the DerSimonian and Laird procedure) and plotted  
178 against mean BMI value within each category. To make comparisons across BMI groups  
179 irrespective of the choice of a reference group, a floating variance estimate (reflecting independent  
180 variability within each group, including the reference group) was attributed to each category using  
181 Plummer's method and used to calculate group-specific confidence intervals (CIs).<sup>20</sup>

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183 To estimate BMI levels at which mortality risk was lowest (i.e., the nadir), weighted linear  
184 regression yielded the best-fitting second-degree fractional polynomial model relating pooled log  
185 HRs to pooled mean BMI levels (weighted by the inverse of the floating variance of the log HR),  
186 and the minimum of this polynomial was the nadir.

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We assessed all-cause mortality and its main components, coronary heart disease, stroke, cardiovascular disease, cancer, and respiratory disease (**eTable 3**). HRs were calculated separately within each study using Cox regression models stratified for age and sex (**eAppendix 2**). Attained age was used, with participants contributing from the baseline survey in crude analyses or from year 5 in the primary analyses. HRs in sex- and age-specific groups (and, where appropriate, by trial arms) were combined across studies.<sup>21</sup> To avoid over-fitting of statistical models, studies with  $\leq 10$  deaths from a particular cause were excluded from meta-analyses of that cause.<sup>22,23</sup>

Because BMI-mortality associations were approximately log-linear above a BMI of 25 kg/m<sup>2</sup>, we calculated HRs per 5 kg/m<sup>2</sup> increase by inverse-variance-weighted regression of the pooled log HRs on mean BMI values in each category.<sup>16</sup> For all-cause mortality we estimate population-attributable fractions (PAFs) for underweight, overweight and obese by combining the proportional excess mortality versus normal weight (HR-1) in these BMI categories ( $X_0$ ,  $X_1$  and  $X_2$  respectively) with the corresponding prevalences ( $P_0$ ,  $P_1$  and  $P_2$ , taken from Global Burden of Disease<sup>24</sup> region-specific prevalences). The PAFs for overweight and obesity are then  $P_1X_1/k$  and  $P_2X_2/k$ , where  $k=1+P_0X_0+P_1X_1+P_2X_2$ . Between-study heterogeneity was quantified by the  $I^2$  statistic.<sup>25</sup> We used 2-sided P-values and 95% CIs.

**Role of funders**

No sponsor had any role in study design, conduct analysis or interpretation; in manuscript preparation, review, or approval; or in deciding to submit for publication. SK, PG, EDA, and JD had access to all the data, and, together with SNB and FBH, were responsible for the decision to submit the manuscript.

## RESULTS

Of 10,625,411 participants from 239 studies (median follow-up 13.7 [IQR 11.4-14.7] years), 3,951,455 were never-smokers without specific chronic diseases at recruitment who survived 5 years, of whom 385,879 died. To limit bias, the pre-specified primary analyses involve this restricted population. To avoid merging importantly different risks, many of these primary analyses further subdivide the WHO-defined normal and overweight BMI categories, yielding 9 BMI groups rather than 6.

**Table 1** shows the substantial relevance of successively stricter exclusions, going from crude analyses of 10.6 million to pre-specified analyses of 4 million adults. With BMI in only 6 groups the whole range from 18.5 to <25 kg/m<sup>2</sup> is the reference group, and HRs were: underweight 1.47 (95% CI 1.39-1.55), overweight 1.11 (1.10-1.11), and grade I, II, III obesity 1.44 (1.41-1.47), 1.92 (1.86-1.98), and 2.71 (2.55-2.86), respectively. With normal and overweight more finely subdivided, however, BMI 22.5-<25 becomes the reference group, and mortality was minimal at BMI 20-<25, and was significantly increased just below this range (BMI 18.5-<20: 1.13, 1.09-1.17; 15-<18.5: 1.51, 1.43-1.59) and *throughout* the overweight range just above it (BMI 25-<27.5: HR=1.07, 1.07-1.08; BMI 27.5-<30: HR=1.20, 1.18-1.22). With this more precise reference group, the HRs for grade I, II, III obesity increased to 1.45 (1.41-1.48), 1.94 (1.87-2.01), and 2.76 (2.60-2.92).

In these pre-specified analyses of 4 million adults, the HRs for overweight and for obesity grade I were broadly similar across different geographic regions (Europe, North America, East Asia, and Australia/NZ; numbers of deaths in South Asia were too small to be reliable), but the HRs for underweight and grade III obesity appeared somewhat higher in Europe than in East Asia (**Figure 1, Table 2, eTables 4-9**).

Combining all regions, the HRs for overweight and obesity were higher in men than women, and higher at younger than older ages (**Figures 2-3**); this held in each major geographic region (**eFigures 2-4, eTables 10-11**). In each region, BMI was non-linearly associated with all-cause mortality, with nadir at BMI 20-<25 and excess mortality in underweight, overweight, and at BMI



241 18.5-<20, at the lower end of the WHO-defined normal range. The nadir depended on age: it was  
242 BMI=22 for age 35-49 years, BMI=23 for 50-69 years, and BMI=24 for 70-89 years).

243  
244 Population-attributable fractions for all-cause mortality due to overweight or obesity were 18%  
245 North America, 15% Europe, 10% Australia/NZ, but only 5% East Asia (**eTable 12**). For BMI $\geq$ 25,  
246 the association of BMI with all-cause mortality was approximately log-linear, and of similar  
247 strength in each region (except perhaps South Asia, where numbers were limited), with HR per 5  
248 units higher BMI 1.31 (1.29-1.33) overall, 1.39 (1.34-1.44) in East Asia, 1.39 (1.34-1.43) in  
249 Europe, 1.29 (1.26-1.32) in North America, and 1.31 (1.27, 1.35) in Australia/NZ. It was 1.51  
250 (1.46-1.56) for men as against only 1.30 (1.26-1.33) for women, heterogeneity  $P<0.0001$ , and it  
251 decreased with age from 1.52 (1.47-1.56) for ages 35-49 years at baseline to 1.21 (1.17-1.25) for  
252 ages 70-89 years, trend  $P<0.0001$  (**eTable 13**).

253  
254 For each major cause of death, BMI was non-linearly associated with mortality in each major  
255 region we studied (**Figure 4, eFigure 5, eTables 14-15**). Above 25 kg/m<sup>2</sup>, it was strongly  
256 positively related to coronary, stroke and respiratory mortality, and moderately related to cancer  
257 mortality. Findings for overweight and obesity were broadly similar in Europe, North America and  
258 East Asia. Within the WHO's wide normal BMI range (18.5-<25 kg/m<sup>2</sup>) the main geographic  
259 difference was that in East Asia mortality from coronary heart disease was steeply lower with  
260 decreasing BMI (having nadir at 18.5-<20 kg/m<sup>2</sup>), but in other regions it was not (**eTable 15**). In  
261 all regions, underweight was associated with substantially higher respiratory mortality and  
262 somewhat higher mortality from coronary heart disease, stroke, and cancer. HRs comparing  
263 underweight versus normal weight cardiovascular mortality were more extreme in Europe than  
264 elsewhere.

265  
266 Compared with the strict primary analyses noted above, crude analyses that ignored smoking,  
267 ignored any effects of prior disease at baseline and failed to exclude the first 5 years of follow-up  
268 yielded different (presumably substantially biased) results, with exaggerated HRs for underweight,  
269 no apparent HRs for overweight and less than half of the less-biased HRs for grade I obesity  
270 (**Table 1, eFigure 6**). In sensitivity analyses (**eFigures 7-11, eTables 17-23**), HRs were little

271 changed in analyses that: used fixed effect models or restricted follow-up to years 5-15;  
272 considered age at risk rather than age at baseline; used contemporary mortality rates; adjusted  
273 additionally for race or excluded participants with diabetes at baseline; used only studies that  
274 included both sexes; used only studies with baseline data on heart disease, stroke and cancer; or  
275 subdivided studies by mean baseline BMI or median recruitment year (HRs were somewhat higher  
276 in studies starting before than since 1990, but meta-regression of HRs on year of recruitment was  
277 not significant). HRs did not vary substantially between larger and smaller studies, between studies  
278 with measured and self-reported BMI, or between occupational and other studies.

## 280 **DISCUSSION**

281 Associations between BMI and mortality can help estimate the public health impact of excess  
282 adiposity only if the estimated relationships are not substantially distorted by the effects of  
283 smoking or ill health on BMI. Hence, our primary analyses were of non-smokers without prior  
284 disease who survived at least 5 years. We conducted standardised comparison of associations of  
285 BMI with mortality across prospective studies in four continents. Both overweight and obesity were  
286 associated with all-cause mortality. In the BMI range above 25 kg/m<sup>2</sup> (the upper limit of the WHO's  
287 normal range) the relationship to mortality was steep in every global region we studied, except  
288 perhaps South Asia where numbers were small.<sup>26</sup>

289  
290 Our primary analyses challenge previous suggestions that overweight (25-<30) and grade I  
291 obesity (30-<35) are not associated with higher mortality,<sup>27</sup> bypassing speculation about  
292 hypothetical protective metabolic effects of increased body fat in apparently healthy individuals.<sup>28</sup>  
293 In particular, the findings here contrast with those of a recent review of published data which  
294 found that, relative to normal weight, grade I obesity was not associated with higher mortality and  
295 overweight was associated with lower all-cause mortality.<sup>27</sup> That review could not, however,  
296 control for the biases controlled for in the current analysis. Indeed, the results of the current  
297 analysis (eg, Table 1 and eTables 4-5) show how that literature-based review's limited ability to  
298 control for bias could have accounted for its misleading findings. Our study was able to reproduce  
299 such findings when conducting crude analyses with inadequate control of reverse causality, but not  
300 when we conducted appropriately strict analyses.

301 Despite broadly similar overall findings across different continents, we found some differences. HRs  
302 per 5-unit increment of BMI above 25 kg/m<sup>2</sup> were higher in Europe than in North America. HRs  
303 were higher in males than females, consistent with previous observations that, at equivalent BMI  
304 levels, men have greater insulin resistance, ectopic (e.g., liver) fat levels and type 2 diabetes  
305 prevalence.<sup>29</sup> In each major region we studied, HRs were substantially higher at younger than at  
306 older ages, although the absolute excess mortality rate was higher in older people.

307  
308 Partly because the prevalence of obesity differs by region, for all-cause mortality there was wide  
309 variation across regions in the approximate population-attributable fraction due to overweight and  
310 obesity: North America 18%, Europe 15%, Australia/New Zealand 10%, and East Asia 5%. These  
311 findings suggest that if the entire population had ideal levels of BMI this would avoid about 1 in 5  
312 premature deaths in North America, 1 in 7 in Europe, 1 in 10 in Australia/New Zealand, and 1 in 20  
313 in East Asia, assuming that the associations of overweight and obesity with mortality in our  
314 primary analyses largely reflect causal effects. Instead, however, BMI is increasing in many  
315 populations, so the pattern of high mortality from adiposity in North America may become typical  
316 elsewhere unless this increase in overweight and obesity can be halted.<sup>30</sup> At the opposite extreme,  
317 there was a substantially higher mortality rate not only among those in WHO's underweight  
318 category, but also in those with BMI 18.5-<20 kg/m<sup>2</sup>, suggesting that in excessively lean  
319 populations underweight remains a cause for concern. We have no information on whether the BMI  
320 in underweight individuals was always low.

321  
322 Our primary analyses used three main approaches to help avoid bias. First, we restricted analysis  
323 to never-smokers to avoid as fully as possible residual confounding by smoking because merely  
324 adjusting for smoking habits would be unlikely to eliminate important residual biases due to the  
325 effect on BMI of different intensity of smoking.<sup>12</sup> Second, we excluded people known to have  
326 certain pre-existing chronic diseases (although full information on this was often unavailable).  
327 Finally, we omitted the initial five years of follow-up from the analysis because conditions at  
328 baseline that might cause death over the next 5 years, could result in reverse causation (where  
329 lower BMI at recruitment is the result, rather than the cause, of the underlying pathology).<sup>13-15</sup>

330 Our findings are consistent with other (albeit less precise) studies that have used effective  
331 methods to reduce potential bias in evaluations of a causal relationship between excess BMI and  
332 mortality, such as Mendelian randomisation analyses,<sup>31,32</sup> other instrumental variable analyses,<sup>33</sup>  
333 and a meta-analysis of randomised trials.<sup>34</sup> Our findings are also broadly consistent with the  
334 stricter analyses done in a recent study of 12 million Korean adults.<sup>35</sup>

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336 The most important limitation is that our only measure of adiposity was BMI, so we could not  
337 address aspects of body composition such as visceral fat or fat distribution,<sup>36,37</sup> nor could we  
338 consider modification of HRs by metabolic factors.<sup>38</sup> Such factors may have different effects in  
339 different populations as, given the same BMI, people of Asian ancestry may have higher amounts  
340 of body fat and greater risk of developing metabolic diseases than people of European ancestry.<sup>39</sup>  
341 Moreover, South Asia, Africa, and Latin America were either unrepresented or poorly represented,  
342 and large studies in those areas might yield somewhat different findings. The study-specific results  
343 were in general not adjusted for ethnicity, or for socioeconomic status. We did not adjust for  
344 regression dilution because previous surveys have reported high levels of concordance in replicate  
345 BMI measures taken from the same adults some years apart.<sup>40</sup>

346  
347 There are, however, particular strengths. Compared with single-country studies, we enhanced  
348 generalisability by combining findings from 239 studies across four continents. We had access to  
349 data for about 97% of the participants in the studies eligible for this analysis (giving large numbers  
350 and negligible bias from unavailability of particular studies), we used a pre-specified analysis plan,  
351 we analysed individual-participant data to avoid the potentially important limitations of literature-  
352 based reviews,<sup>41</sup> we analysed clinically relevant subpopulations reliably, exploiting the considerable  
353 statistical power of the study. We avoided potential “over-adjustment” by not adjusting for  
354 variables (e.g, diabetes status, physical activity) that could mediate associations between BMI and  
355 mortality.<sup>42</sup> Finally, our results were robust to a variety of sensitivity analyses.

356  
357 We conclude that wherever overweight and obesity are common their associations with higher all-  
358 cause mortality are positive and broadly similar, supporting strategies to combat the entire  
359 spectrum of excessive adiposity worldwide.

360 **Dedication:** This paper is dedicated to the memory of Gary Whitlock, who contributed much to  
361 developing this collaboration.

362  
363 **Funding:** <http://www.phpc.cam.ac.uk/ceu/research/global-bmi-mortality-collaboration/> provides  
364 links to websites of the component studies (or consortia), many of which describe their funding.  
365 The coordinating centre at the University of Cambridge was funded by the UK Medical Research  
366 Council (G0800270), British Heart Foundation (SP/09/002), British Heart Foundation Cambridge  
367 Cardiovascular Centre of Excellence, and UK National Institute for Health Research Cambridge  
368 Biomedical Research Centre. The work of the coordinating center at the Harvard T.H. Chan School  
369 of Public Health was funded by grants P01 CA87969, UM1 CA176726, UM1 CA167552, DK58845,  
370 P30 DK046200, and U54 CA155626 from the National Institutes of Health. This research has been  
371 conducted using the UK Biobank resource.

372  
373 **Contributions:** All of the authors contributed to data collection, and the design, analysis,  
374 interpretation, and re-drafting of this paper. SK, PG, and DW conducted the combined statistical  
375 analysis. EDA, SNB, JD, and FBH drafted the manuscript.

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## References

1. Ng M, Fleming T, Robinson M et al. Global, regional, and national prevalence of overweight and obesity in children and adults during 1980–2013: a systematic analysis for the Global Burden of Disease Study 2013. *Lancet*. 2014;384(9945):766-781.
2. World Health Organization. Obesity and overweight. *Fact sheet N°311 (Updated January 2015)*. 2015.
3. WHO expert consultation. Appropriate body-mass index for Asian populations and its implications for policy and intervention strategies. *Lancet*. 2004;363(9403):157-163.
4. Asia Pacific Cohort Studies Collaboration. Body mass index and cardiovascular disease in the Asia-Pacific Region: an overview of 33 cohorts involving 310 000 participants. *International Journal of Epidemiology*. 2004;33(4):751-758.
5. Berrington de Gonzalez A, Hartge P, Cerhan JR et al. Body-Mass Index and Mortality among 1.46 Million White Adults. *N Engl J Med*. 2010;363(23):2211-2219.
6. Jee SH, Sull JW, Park J et al. Body-Mass Index and Mortality in Korean Men and Women. *N Engl J Med*. 2006;355(8):779-787.
7. Patel AV, Hildebrand JS, Gapstur SM. Body Mass Index and All-Cause Mortality in a Large Prospective Cohort of White and Black U.S. Adults. *PLoS ONE*. 2014;9(10):e109153.
8. Pischon T, Boeing H, Hoffmann K et al. General and Abdominal Adiposity and Risk of Death in Europe. *N Engl J Med*. 2008;359(20):2105-2120.
9. Prospective Studies Collaboration. Body-mass index and cause-specific mortality in 900000 adults: collaborative analyses of 57 prospective studies. *Lancet*. 2009;373(9669):1083-1096.
10. The Emerging Risk Factors Collaboration. Separate and combined associations of body-mass index and abdominal adiposity with cardiovascular disease: collaborative analysis of 58 prospective studies. *Lancet*. 2011;377(9771):1085-1095.
11. Zheng W, McLerran DF, Rolland B et al. Association between Body-Mass Index and Risk of Death in More Than 1 Million Asians. *N Engl J Med*. 2011;364(8):719-729.
12. Bamia C, Trichopoulou A, Lenas D, Trichopoulos D. Tobacco smoking in relation to body fat mass and distribution in a general population sample. *Int J Obes Relat Metab Disord*. 2004;28(8):1091-1096.
13. Manson JE, Stampfer MJ, Hennekens CH, Willett WC. Body weight and longevity: A reassessment. *JAMA*. 1987;257(3):353-358.
14. Singh PN, Wang X. Simulation Study of the Effect of the Early Mortality Exclusion on Confounding of the Exposure-Mortality Relation by Preexisting Disease. *American Journal of Epidemiology*. 2001;154(10):963-971.
15. Willett WC, Hu FB, Thun M. Overweight, obesity, and all-cause mortality. *JAMA*. 2013;309(16):1681-1682.
16. Stewart LA, Clarke M, Rovers M. Preferred reporting items for a systematic review and meta-analysis of individual participant data: The prisma-ipd statement. *JAMA*. 2015;313(16):1657-1665.
17. Park SY, Wilkens L, Murphy S, Monroe K, Henderson B, Kolonel L. Body mass index and mortality in an ethnically diverse population: the Multiethnic Cohort Study. *Eur J Epidemiol*. 2012;27(7):489-497.

- 419 18. Lin WY, Tsai SL, Albu JB et al. Body mass index and all-cause mortality in a large Chinese  
420 cohort. *Canadian Medical Association Journal*. 2011;183(6):E329-E336.
- 421 19. Clinical Guidelines on the Identification, Evaluation, and Treatment of Overweight and Obesity  
422 in Adults--The Evidence Report. National Institutes of Health. *Obes Res*. 1998;6 Suppl 2:51S-  
423 209S.
- 424 20. Plummer M. Improved estimates of floating absolute risk. *Statist Med*. 2004;23(1):93-104.
- 425 21. Jackson D, White IR, Thompson SG. Extending DerSimonian and Laird's methodology to  
426 perform multivariate random effects meta-analyses. *Statist Med*. 2010;29(12):1282-1297.
- 427 22. Preston SH, Mehta NK, Stokes A. Modeling Obesity Histories in Cohort Analyses of Health and  
428 Mortality. *Epidemiology*. 2013;24(1):158-66.
- 429 23. Thompson S, Kaptoge S, White I, Wood A, Perry P, Danesh J. Statistical methods for the  
430 time-to-event analysis of individual participant data from multiple epidemiological studies. *Int  
431 J Epidemiol*. 2010;39(5):1345-1359.
- 432 24. Stevens G, Singh G, Lu Y et al. National, regional, and global trends in adult overweight and  
433 obesity prevalences. *Population Health Metrics*. 2012;10(1):22.
- 434 25. Higgins JP, Thompson SG. Quantifying heterogeneity in a meta-analysis. *Stat Med*.  
435 2002;21(11):1539-1558.
- 436 26. Jensen MD, Ryan DH, Apovian CM et al. 2013 AHA/ACC/TOS Guideline for the Management of  
437 Overweight and Obesity in Adults: A Report of the American College of Cardiology/American  
438 Heart Association Task Force on Practice Guidelines and The Obesity Society. *Circulation*.  
439 2014;129(25 suppl 2):S102-S138.
- 440 27. Flegal KM, Kit BK, Orpana H, Graubard BI. Association of all-cause mortality with overweight  
441 and obesity using standard body mass index categories: A systematic review and meta-  
442 analysis. *JAMA*. 2013;309(1):71-82.
- 443 28. Doehner W, Clark A, Anker SD. The obesity paradox: weighing the benefit. *European Heart  
444 Journal*. 2010;31(2):146-8.
- 445 29. Sattar N. Gender aspects in type 2 diabetes mellitus and cardiometabolic risk. *Best Practice &  
446 Research Clinical Endocrinology & Metabolism*. 2013;27(4):501-507.
- 447 30. Gillman MW, Ludwig DS. How Early Should Obesity Prevention Start? *N Engl J Med*.  
448 2013;369(23):2173-2175.
- 449 31. Nordestgaard BG, Palmer TM, Benn M et al. The effect of elevated body mass index on  
450 ischemic heart disease risk: causal estimates from a Mendelian randomisation approach. *PLoS  
451 Med*. 2012;9(5):e1001212.
- 452 32. Hagg S, Fall T, Ploner A et al. Adiposity as a cause of cardiovascular disease: a Mendelian  
453 randomization study. *International Journal of Epidemiology*. 2015;44(2):578-586.
- 454 33. Davey Smith G, Sterne JA, Fraser A, Tynelius P, Lawlor DA, Rasmussen F. The association  
455 between BMI and mortality using offspring BMI as an indicator of own BMI: large  
456 intergenerational mortality study. *BMJ*. 2009;339:b5043.
- 457 34. Kritchevsky SB, Beavers KM, Miller ME et al. Intentional Weight Loss and All-Cause Mortality:  
458 A Meta-Analysis of Randomized Clinical Trials. *PLoS ONE*. 2015;10(3):e0121993.

- 459 35. Yi SW, Ohrr H, Shin SA, Yi JJ. Sex-age-specific association of body mass index with all-cause  
460 mortality among 12.8 million Korean adults: a prospective cohort study. *International Journal*  
461 *of Epidemiology*. 2015;44(5):1696-1705.
- 462 36. Bohm A, Heitmann BL. The use of bioelectrical impedance analysis for body composition in  
463 epidemiological studies. *Eur J Clin Nutr*. 2013;67(S1):S79-S85.
- 464 37. Britton KA, Massaro JM, Murabito JM, Kreger BE, Hoffmann U, Fox CS. Body fat distribution,  
465 incident cardiovascular disease, cancer, and all-cause mortality. *J Am Coll Cardiol*.  
466 2013;62(10):921-925.
- 467 38. Kramer CK, Zinman B, Retnakaran R. Are Metabolically Healthy Overweight and Obesity  
468 Benign Conditions? A Systematic Review and Meta-analysis. *Annals of Internal Medicine*.  
469 2013;159(11):758-769.
- 470 39. Chan JN, Malik V, Jia W. Diabetes in asia: Epidemiology, risk factors, and pathophysiology.  
471 *JAMA*. 2009;301(20):2129-2140.
- 472 40. Wormser D, White IR, Thompson SG, Wood AM. Within-person variability in calculated risk  
473 factors: Comparing the aetiological association of adiposity ratios with risk of coronary heart  
474 disease. *International Journal of Epidemiology*. 2013;42(3):849-859.
- 475 41. Stewart LA, Clarke MJ. Practical methodology of meta-analyses (overviews) using updated  
476 individual patient data. Cochrane Working Group. *Stat Med*. 1995;14(19):2057-2079.
- 477 42. Van Gaal LF, Mertens IL, De Block CE. Mechanisms linking obesity with cardiovascular  
478 disease. *Nature*. 2006;444(7121):875-880.  
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**Table 1: Effects of successively stricter precautions against bias on analyses of BMI vs all-cause mortality**

	<b>Underweight (BMI 15 to &lt;18.5)</b>	<b>Normal weight (18.5 to &lt;25)</b>	<b>Overweight (25 to &lt;30)</b>	<b>Obesity Grade I (30 to &lt;35)</b>	<b>Obesity Grade II (35 to &lt;40)</b>	<b>Obesity Grade III (40 to &lt;60)</b>			
<b>6 WHO BMI groups: Crude analysis with no exclusions (237 studies; 10,622,450 participants; 1,601,774 deaths)</b>									
Participants / deaths	292003 / 68455	5586892 / 810838	3467617 / 526098	946257 / 144871	237223 / 36113	92458 / 15399			
HR (95% CI)	1.82 (1.74, 1.91)	1.00 (0.98, 1.02)	0.95 (0.94, 0.97)	1.17 (1.16, 1.18)	1.49 (1.47, 1.51)	1.95 (1.90, 2.01)			
<b>6 WHO BMI groups: Participants without known disease at baseline (236 studies; 9,104,247 participants; 1,210,250 deaths)</b>									
Participants / deaths	255000 / 52789	4922817 / 631488	2916978 / 388781	756075 / 102315	183689 / 24556	69688 / 10321			
HR (95% CI)	1.81 (1.72, 1.91)	1.00 (0.98, 1.02)	0.95 (0.95, 0.96)	1.18 (1.16, 1.20)	1.52 (1.48, 1.55)	2.05 (1.98, 2.13)			
<b>6 WHO BMI groups: Participants without known chronic disease at baseline, adjusting for smoking status (234 studies; 8,801,617 participants; 1,185,728 deaths)</b>									
Participants / deaths	245080 / 51170	4751019 / 618881	2826687 / 381617	733108 / 100113	178130 / 23945	67593 / 10002			
HR (95% CI)	1.70 (1.61, 1.80)	1.00 (0.98, 1.02)	0.99 (0.98, 1.00)	1.25 (1.23, 1.27)	1.63 (1.59, 1.66)	2.24 (2.15, 2.33)			
<b>6 WHO BMI groups: Participants without known chronic disease at baseline, adjusting for smoking status, and excluding the first 5 years of follow-up (213 studies; 7,805,434 participants; 949,010 deaths)</b>									
Participants / deaths	208044 / 33817	4234052 / 496310	2513128 / 312450	641237 / 80037	152741 / 18737	56232 / 7659			
HR (95% CI)	1.60 (1.51, 1.70)	1.00 (0.98, 1.02)	1.03 (1.01, 1.04)	1.31 (1.29, 1.33)	1.70 (1.67, 1.74)	2.36 (2.27, 2.45)			
<b>6 WHO BMI groups (a primary pre-specified analysis): Never-smokers without known chronic disease at baseline, excluding the first 5 years of follow-up (189 studies; 3,951,455 participants; 385,879 deaths)</b>									
Participants / deaths	114091 / 12726	2145550 / 192523	1250103 / 130293	330840 / 37318	80827 / 9179	30044 / 3840			
HR (95% CI)	1.47 (1.39, 1.55)	1.00 (0.98, 1.02)	1.11 (1.10, 1.11)	1.44 (1.41, 1.47)	1.92 (1.86, 1.98)	2.71 (2.55, 2.86)			
<b>9 BMI groups (strictest precautions against bias: a primary pre-specified analysis): As above, but with normal weight and overweight further subdivided (189 studies; 3,951,455 participants; 385,879 deaths)</b>									
<b>BMI categories</b>	<b>15 to &lt;18.5</b>	<b>18.5 to &lt;20</b>	<b>20 to &lt;22.5</b>	<b>22.5 to &lt;25</b>	<b>25 to &lt;27.5</b>	<b>27.5 to &lt;30</b>	<b>30 to &lt;35</b>	<b>35 to &lt;40</b>	<b>40 to &lt;60</b>
Participants/deaths	114091/12726	230749/20989	838907/72701	1075894/98833	821303/84952	428800/45341	330840/37318	80827/9179	30044/3840
HR (95% CI)	1.51 (1.43, 1.59)	1.13 (1.09, 1.17)	1.00 (0.98, 1.02)	1.00 (0.99, 1.01)	1.07 (1.07, 1.08)	1.20 (1.18, 1.22)	1.45 (1.41, 1.48)	1.94 (1.87, 2.01)	2.76 (2.60, 2.92)

All analyses are adjusted for age and sex

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**Table 2: Strictest pre-specified analyses, in geographic regions with >1 million participants: Nine BMI groups vs all-cause mortality among never-smokers, excluding chronic disease at baseline and 5 years follow-up**

<b>BMI categories</b>	<b>15 to &lt;18.5</b>	<b>18.5 to &lt;20</b>	<b>20 to &lt;22.5</b>	<b>22.5 to &lt;25</b>	<b>25 to &lt;27.5</b>	<b>27.5 to &lt;30</b>	<b>30 to &lt;35.0</b>	<b>35 to &lt;40</b>	<b>40 to &lt;60</b>
<b>Europe (89 studies; 1,135,600 participants; 56,477 deaths)</b>									
Participants/deaths	13398/675	42584/1508	199369/7449	306566/13278	249929/12850	153147/8935	127536/8386	32749/2424	10322/972
HR (95% CI)	1.79 (1.63, 1.97)	1.25 (1.14, 1.38)	1.02 (0.97, 1.07)	1.00 (0.97, 1.03)	1.07 (1.06, 1.09)	1.21 (1.18, 1.25)	1.52 (1.45, 1.58)	1.99 (1.87, 2.12)	3.04 (2.84, 3.27)
<b>North America (40 studies; 1,415,087 participants; 219,922 deaths)</b>									
Participants/deaths	22028/3846	67114/8597	274883/36200	359022/54995	317721/53464	168183/28471	149807/25348	39379/6299	16950/2702
HR (95% CI)	1.51 (1.34, 1.70)	1.09 (1.02, 1.16)	1.01 (0.96, 1.06)	1.00 (0.97, 1.03)	1.06 (1.04, 1.07)	1.17 (1.12, 1.22)	1.39 (1.30, 1.49)	1.93 (1.74, 2.13)	2.58 (2.26, 2.93)
<b>East Asia (46 studies; 1,074,385 participants; 100,784 deaths)</b>									
Participants/deaths	46979/7178	94409/10206	301242/27537	336758/28755	194857/17070	72133/6950	25658/2753	1941/231	408/104
HR (95% CI)	1.36 (1.25, 1.49)	1.11 (1.04, 1.18)	0.99 (0.97, 1.02)	1.00 (0.97, 1.03)	1.07 (1.04, 1.11)	1.28 (1.21, 1.35)	1.54 (1.42, 1.67)	2.01 (1.59, 2.54)	2.38 (1.33, 4.24)
P-value for heterogeneity	0.0045	0.28	0.42	-	0.89	0.46	0.20	0.48	<0.0001

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Normal weight and overweight are subdivided, and the reference category is BMI 22.5 to <25 kg/m<sup>2</sup>.

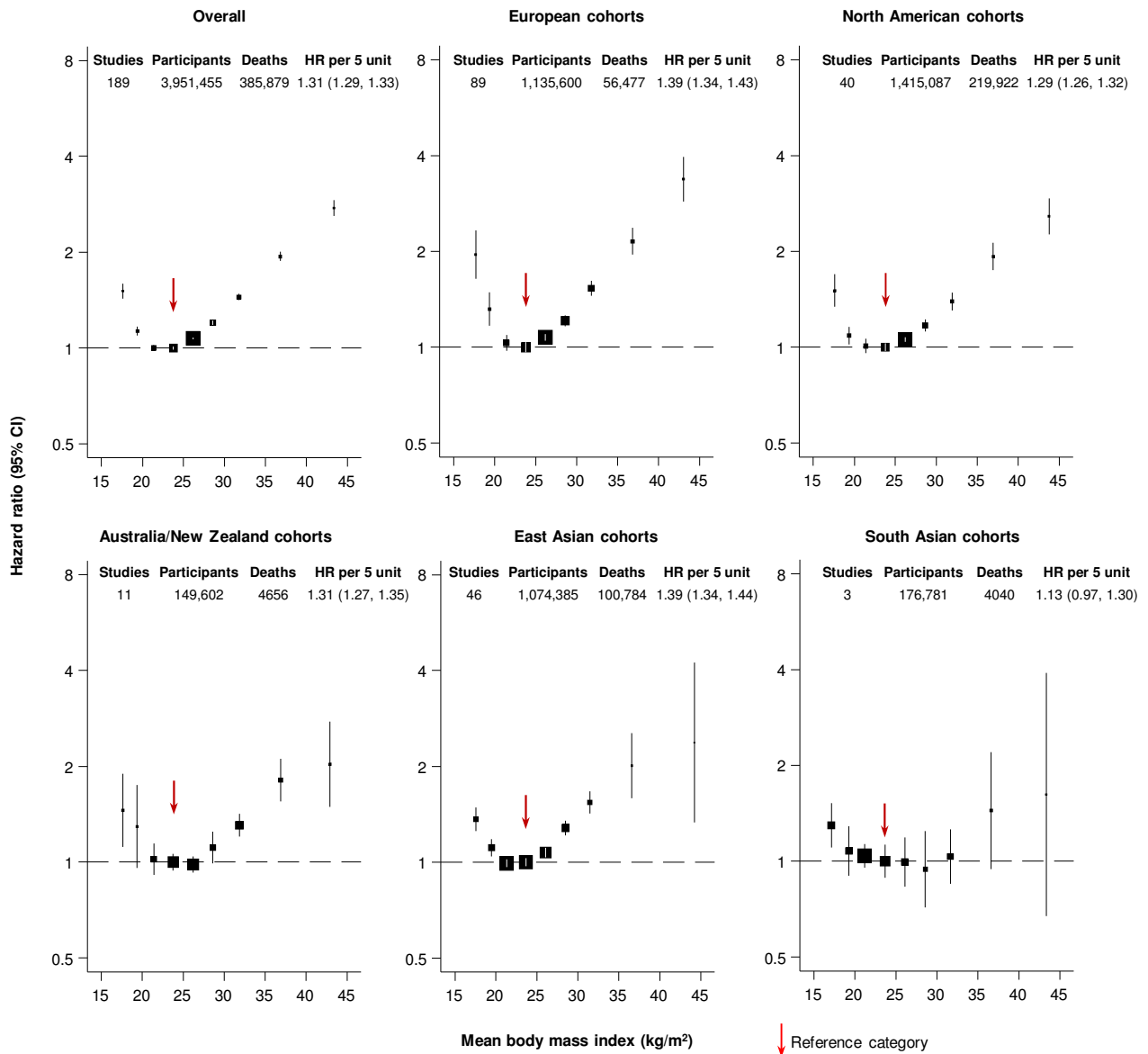
Numbers of studies, participants and deaths are after exclusions from these pre-specified principal analyses.

CI's were calculated using floating variance estimates (reflecting independent variability within each group, including the reference group).

Results from studies in South Asia and Australia/New Zealand are in Figure 1, with details in eTable 8 of the Supplementary Material.

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**Figure 1:** Association of BMI with all-cause mortality, by geographical region



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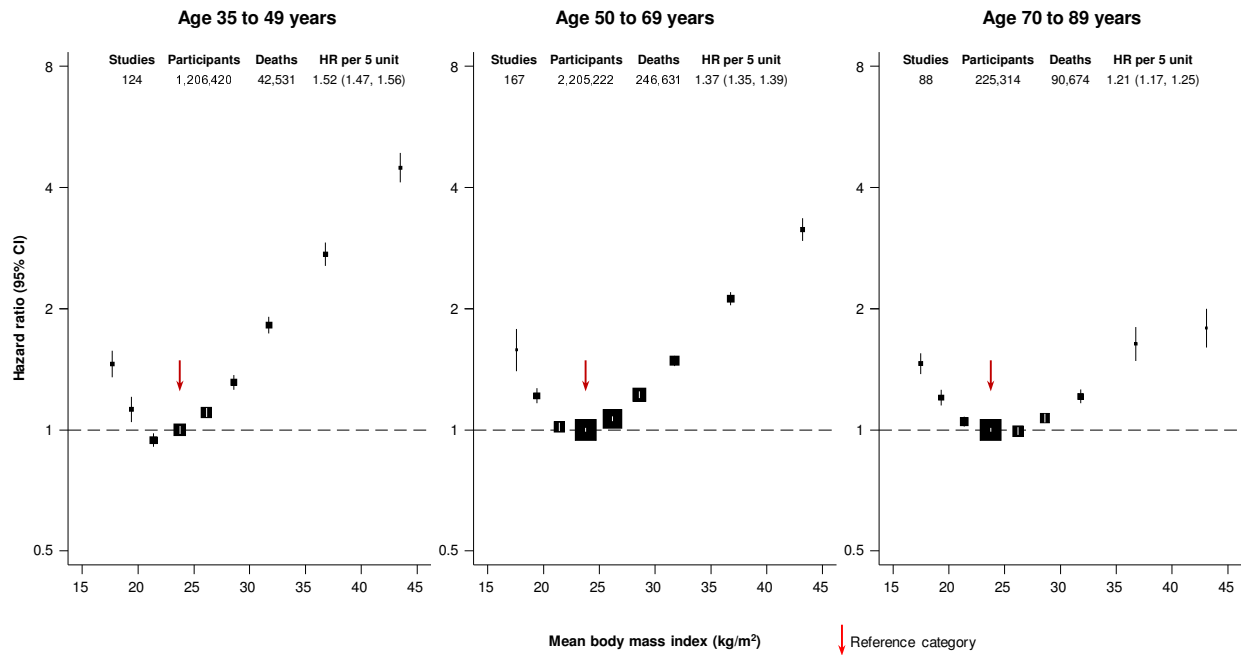
HR, hazard ratio per 5 kg/m<sup>2</sup> unit of BMI in the range BMI >25 kg/m<sup>2</sup> (and 95% CI).

Analyses restricted to never-smokers without pre-existing chronic disease, and excluding the first 5 years of follow-up, and include data from all geographical regions.

BMI groups: 15.0 to <18.5, 18.5 to <20, 20 to <22.5, 22.5 to <25.0, 25.0 to <27.5, 27.5 to <30, 30 to <35.0, 35.0 to <40, and 40 to <60 kg/m<sup>2</sup>. Reference category (arrow) is 22.5 to <25 kg/m<sup>2</sup>.

CI's are from floating variance estimates (reflecting independent variability within each category, including reference). Areas of squares are proportional to the information content.

511 **Figure 2:** Association of BMI with all-cause mortality, by baseline age group



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514 HR, hazard ratio per 5 kg/m<sup>2</sup> unit of BMI in the range BMI >25 kg/m<sup>2</sup> (and 95% CI).

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516 Analyses restricted to never-smokers without pre-existing chronic disease, and excluding the first 5 years of follow-up, and include data from all geographical regions.

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518 BMI groups: 15.0 to <18.5, 18.5 to <20, 20 to <22.5, 22.5 to <25.0, 25.0 to <27.5, 27.5 to <30, 30 to <35.0, 35.0 to <40, and 40 to <60 kg/m<sup>2</sup>. Reference category (arrow) is 22.5 to <25 kg/m<sup>2</sup>.

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522 CIs are from floating variance estimates (reflecting independent variability within each category, including reference). Areas of squares are proportional to the information content.

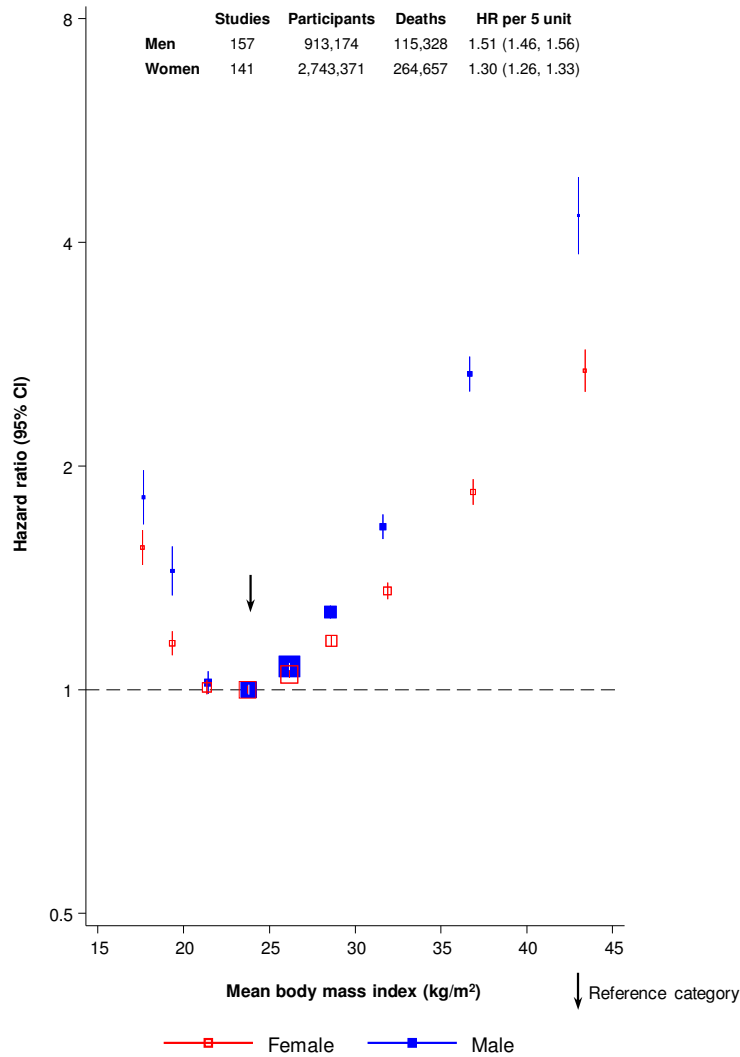
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524 Analyses by age and the 3 main geographic regions are in eFigure 2.

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527 **Figure 3:** Association of BMI with all-cause mortality, by sex



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530 HR, hazard ratio per 5 kg/m<sup>2</sup> unit of BMI in the range BMI >25 kg/m<sup>2</sup> (and 95% CI).

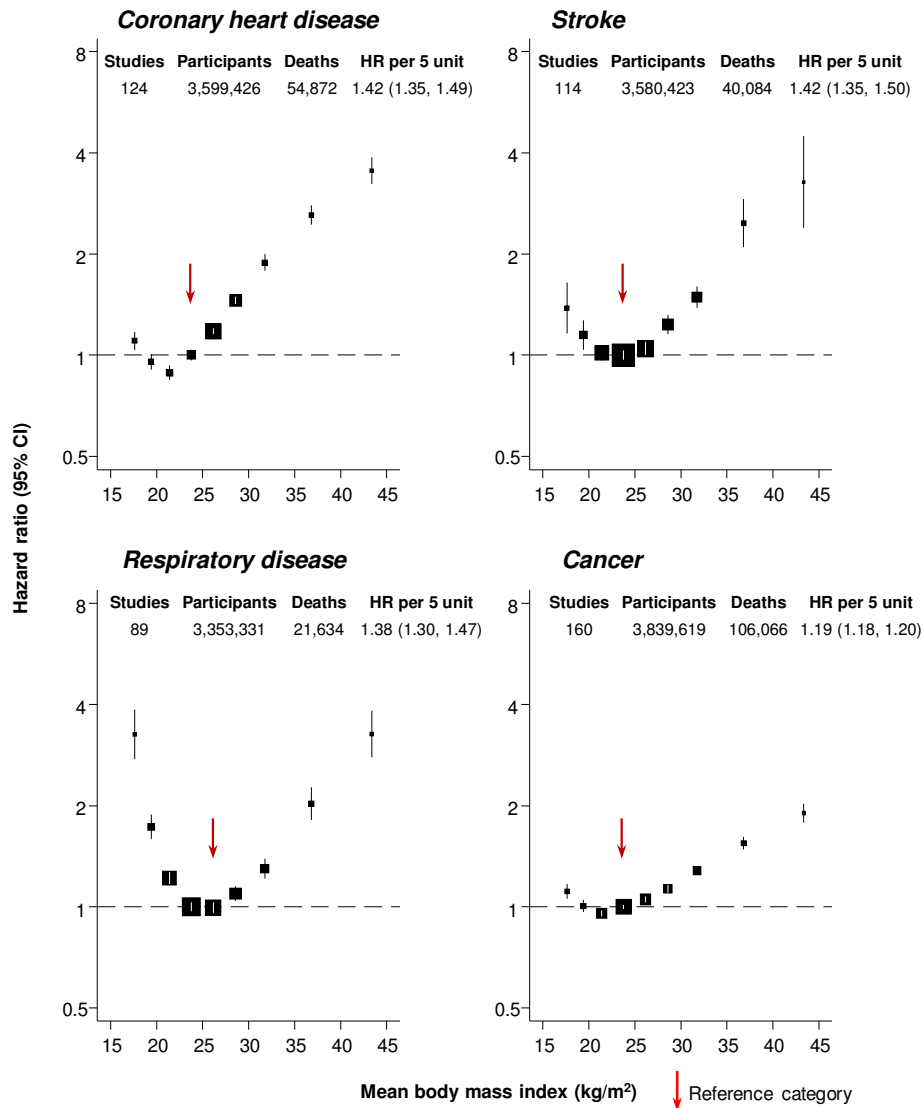
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532 Analyses restricted to never-smokers without pre-existing chronic disease, and excluding the first 5 years of follow-up,  
533 , and include data from all geographical regions.

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535 BMI groups: 15.0 to <18.5, 18.5 to <20, 20 to <22.5, 22.5 to <25.0, 25.0 to <27.5, 27.5 to <30, 30 to <35.0, 35.0  
536 to <40, and 40 to <60 kg/m<sup>2</sup>. Reference category (arrow) is 22.5 to <25 kg/m<sup>2</sup>.

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538 CIs are from floating variance estimates (reflecting independent variability within each category, including reference).  
539 Areas of squares are proportional to the information content.

540 Analyses by sex and the 3 main geographic regions are in eTable 10 and eFigures 3-4.

541 **Figure 4:** Association of BMI with mortality, by major underlying cause



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544 HR, hazard ratio per 5 kg/m<sup>2</sup> unit of BMI in the range BMI >25 kg/m<sup>2</sup> (and 95% CI). Analyses restricted to never-smokers without pre-existing chronic disease, and excluding the first 5 years of follow-up, and include data from all  
 545 geographical regions.  
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547 BMI groups: 15.0 to <18.5, 18.5 to <20, 20 to <22.5, 22.5 to <25.0, 25.0 to <27.5, 27.5 to <30, 30 to <35.0, 35.0  
 548 to <40, and 40 to <60 kg/m<sup>2</sup>. Reference category (arrow) is 22.5 to <25 kg/m<sup>2</sup>.  
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550 CIs are from floating variance estimates (reflecting independent variability within each category, including reference).  
 551 Areas of squares are proportional to the information content.  
 552

553 Analyses of cause-specific mortality by 3 geographic regions are in eTable 15 and eFigure 5.  
 554