

Improving Blood Pressure Control Through Pharmacist Interventions: A Meta-Analysis of Randomized Controlled Trials

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Background—Control of blood pressure (BP) remains a major challenge in primary care. Innovative interventions to improve BP control are therefore needed. By updating and combining data from 2 previous systematic reviews, we assess the effect of pharmacist interventions on BP and identify potential determinants of heterogeneity.

Methods and Results—Randomized controlled trials (RCTs) assessing the effect of pharmacist interventions on BP among outpatients with or without diabetes were identified from MEDLINE, EMBASE, CINAHL, and CENTRAL databases. Weighted mean differences in BP were estimated using random effect models. Prediction intervals (PI) were computed to better express uncertainties in the effect estimates. Thirty-nine RCTs were included with 14 224 patients. Pharmacist interventions mainly included patient education, feedback to physician, and medication management. Compared with usual care, pharmacist interventions showed greater reduction in systolic BP (-7.6 mm Hg, 95% Cl: -9.0 to -6.3; $I^2=67\%$) and diastolic BP (-3.9 mm Hg, 95% Cl: -5.1 to -2.8; $I^2=83\%$). The 95% PI ranged from -13.9 to -1.4 mm Hg for systolic BP and from -9.9 to +2.0 mm Hg for diastolic BP. The effect tended to be larger if the intervention was led by the pharmacist and was done at least monthly.

Conclusions—Pharmacist interventions – alone or in collaboration with other healthcare professionals – improved BP management. Nevertheless, pharmacist interventions had differential effects on BP, from very large to modest or no effect; and determinants of heterogeneity could not be identified. Determining the most efficient, cost-effective, and least time-consuming intervention should be addressed with further research. (*J Am Heart Assoc.* 2014;3:e000718 doi: 10.1161/JAHA.113.000718)

Key Words: hypertension • pharmacist • prediction interval • systematic review • team-based care

E levated blood pressure (BP) is a major cause of death worldwide.¹ Although reduction of BP is a cornerstone of the prevention of cardiovascular diseases (CVD),² numerous hypertensive patients do not achieve adequate BP control. In

Accompanying Table S1 and Figure S1 are available at http://jaha.ahajournals.org/content/3/2/e000718/suppl/DC1

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© 2014 The Authors. Published on behalf of the American Heart Association, Inc., by Wiley Blackwell. This is an open access article under the terms of the Creative Commons Attribution-NonCommercial License, which permits use, distribution and reproduction in any medium, provided the original work is properly cited and is not used for commercial purposes. the United States in 2009-2010, it was estimated that 53% of all hypertensive people and 40% of treated hypertensive people had uncontrolled BP.³ Lower rates of BP control have been reported in European countries.⁴

Innovative interventions to improve BP control are therefore needed in primary care, where management of hypertension takes place. One such approach is a greater use of community-based models of care with the involvement of nonphysician healthcare professionals, specifically, pharmacists and nurses.^{5,6} The US Community Preventive Services Task Force has recently recommended team-based care, including pharmacists, to improve BP control.⁷ Pharmacists are highly accessible healthcare professionals and a valuable asset in the management of hypertension.^{8,9} Recently, we conducted a systematic review and meta-analysis of 30 RCTs underlining that pharmacist interventions - conducted alone or in collaboration with physicians or nurses - substantially reduced systolic and diastolic BP and helped control other major CVD risk factors (total and LDL cholesterol, smoking) among outpatients with CVD risk factors.9 In a second systematic review and meta-analysis of 15 RCTs, we showed

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that pharmacist interventions were associated with a better control of systolic and diastolic BP and other major CVD risk factors (total and LDL cholesterol, body mass index) among outpatients with diabetes.¹⁰

However, in both systematic reviews, a major issue was that various pharmacist interventions were evaluated in different clinical settings with a substantial heterogeneity in the effect estimates on BP. Therefore, some interventions are possibly more efficient than others to reduce BP. The identification of the more efficient pharmacist interventions on BP is of major importance to guide policy and to decide which intervention should be recommended. However, it requires the comparison of data on a large number of studies. Therefore, we combined and updated BP data from these 2 previous systematic reviews to assess the effect of pharmacist interventions on BP among outpatients and to identify determinants of heterogeneity.

Methods

Data were updated and combined from 2 previous systematic reviews of RCTs in which we assessed the effect of pharmacist interventions on CVD risk factors among outpatients.^{9,10} Both reviews were conducted following the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) recommendations¹¹ and using methods outlined in the Cochrane Handbook for Systematic Reviews of Interventions,¹² including the preparation of a protocol and analysis plan.

Research Methods

Details of the research methods are given in the original publications.^{9,10} Briefly, both reviews used the same selection criteria (study design, interventions, and outcomes), search strategy, study selection, data extraction, and assessment of risk of bias across the included studies. The first review included patients with any modifiable CVD risk factors but excluded studies, which targeted patients with diabetes.⁹ The second review included only studies including patients with diabetes.¹⁰ For both reviews, we included RCTs evaluating the effect of pharmacist intervention delivered by a pharmacist alone or in collaboration with other healthcare professionals, among adult outpatients with any modifiable CVD risk factors (hypertension, dyslipidemia, diabetes, smoking, or obesity) compared with the usual care group.

Pharmacist interventions were classified using the a prioridefined categories of pharmacist-directed care (pharmacist initiated and managed interventions) or as pharmacist collaborative care (pharmacist collaborated in interventions conducted with other health care professionals).¹³

Search strategy and study selection

The description of the search strategy and the flow diagram of assessed and included studies in each review are given in detail in the 2 original publications.^{9,10} Briefly, we searched, in collaboration with a research librarian (A.L.C.), the following electronic databases: MEDLINE via PubMed, EMBASE, CI-NAHL, and CENTRAL for published articles using predefined search terms. The first review, which excluded studies on patients with diabetes, retrieved articles up to November 2010, while the second review of patients with diabetes retrieved articles up to March 2012. For the current review, an update of the original literature searches of both reviews was conducted in September 2013. For the current review, we considered only studies where systolic or diastolic BP was the outcome and presented as weighted mean differences in BP.

Data extraction and assessment of risk of bias

A data extraction form was created and pilot tested with some eligible studies and then finalized.

In both reviews, data extraction was independently preformed by 2 authors (V.S. and A.C.). Discrepancies were resolved through discussion. Extracted data included: (1) study author, year of publication, and country where the study was conducted; (2) study characteristics (including study setting and design, duration of follow-up, and sample size); (3) participants characteristics (including sex, age, CVD risk factors, and medications); (4) characteristics of interventions (including description and frequency of the pharmacist intervention); (5) characteristics of usual care group; and (6) types of outcome measure. Details of pharmacist interventions were documented using the a priori-defined categories of pharmacist-directed care or pharmacist collaborative care.¹³

The risk of bias for all included studies was independently assessed by 2 authors (V.S. and A.C.) using the Cochrane Risk of Bias Tool, a validated tool for quality in RCTs,¹² which addresses random sequence generation, allocation concealment, blinding of outcome assessors, completeness of outcome data, selective outcome reporting, and other potential bias (e.g., important baseline imbalance in patient characteristics). Each of these 6 domains in the tool was rated as (1) low risk of bias; (2) unclear risk of bias; and (3) high risk of bias. A study was of high quality if it had a low risk of bias on \geq 4 of the 6 domains in the tool.

Statistical Analysis

Data were analyzed using Stata 12.0 software (StataCorp LP). Outcomes of interest were systolic BP and diastolic BP. Heterogeneity in the effect of interventions was expected across studies and random effect models were used to estimate intervention effects and 95% confidence intervals (Cl).¹⁴ Effects were calculated as weighted mean differences in BP between intervention and usual care group.

Between studies heterogeneity was quantified using the I^2 statistic. To better express uncertainties in the effect estimates, 95% prediction intervals (PI) were computed because they are advocated in case of substantial heterogeneity. While the CI quantifies the accuracy of the mean effect, the PI quantifies the dispersion (or distribution) of effect estimates^{15,16} and will include the true effect from a unique study 95% of the time.¹⁶ The PI is wider than the CI unless there is no heterogeneity.

Reasons for heterogeneity in effect estimates should be sought in meta-analyses.¹⁴ To explore the possible sources of heterogeneity, we conducted subgroup analyses according to selected study characteristics: (1) country where the study was conducted; (2) setting (outpatient clinic versus community pharmacy); (3) including patients with diabetes or not; (4) type of pharmacist care (pharmacist-led care versus collaborative care); (5) type of interventions, (6) including a nurse or not in the intervention; and (7) frequency of intervention. Statistical significance of between subgroup differences was evaluated by meta-regression.¹⁷

Publication bias was evaluated with funnel plots to check symmetrical distribution and convergence toward the pooled effect, along with Egger tests. Further, sensitivity analyses were performed (1) excluding relatively small studies (with fewer than 80 participants per randomization group) and (2) restricting analyses to studies of high quality.

Results

Study Selection

Initially, 30 RCTs with 19 studies having BP outcome were included in the first systematic review.⁹ Of the 5522 additional citations identified in the updated searches of this first review, 33 potentially relevant full-text studies were assessed and screened; 10 additional RCTs met the inclusion criteria, with 6 studies having BP as outcome. Initially, 15 RCTs with 12 studies having BP as outcome were included in the second systematic review.¹⁰ Of the 2742 additional citations identified in the updated searches of the second review, 10 potentially relevant full-text studies were assessed and screened; 2 additional RCTs met the inclusion criteria with 2 studies having BP as outcome. Overall, with the update, data from 39 RCTs with BP as outcome are included in the current review.

Study and Patient Characteristics

Thirty-nine RCTs representing a total of 14 224 patients were included for the current review. Trials were published between

1973 and 2013 in peer-reviewed journals. Details of included studies are given in Table S1. Most trials were conducted in the USA or Canada (N=25); 4 trials¹⁸⁻²¹ were conducted in European countries and 10 trials²²⁻³¹ in Asia, South America, and Australia.

The within-study mean age of participants ranged from 48 to 77 years (mean, 62 years). Overall, 53% of participants were women. The mean duration of follow-up was 8.3 months (median, 6 months; range, 3 to 13 months). For most trials^{19,20,22–28,30–52} (N=31), participants were followed-up in outpatient clinics or by general practitioners; 6 trials^{8,18,19,53–55} were conducted in community pharmacy and 2 trials^{21,29} in both outpatient clinics and community pharmacies.

The interventions were led by the pharmacist in 23 studies and conducted in collaboration with other healthcare professionals, such as physician, nurse, dietitian, or physical therapist, in 16 studies (Table S1). In the 16 pharmacist collaborative care studies, the team composition was pharmacist/physician (N=10); pharmacist/nurse/physician (N=3); pharmacist/nurse/physician/dietitian or nutritionist therapist (N=2), or pharmacist/nurse (N=1). Overall, a nurse was involved in the intervention in 6 studies.

The interventions consisted of (1) patient education and counseling about lifestyle, medication and medication adherence (N=35 studies); (2) feedback to healthcare professional (including drug-related problems identification; recommendation to physician for medication change; team meeting, development of treatment plan) (N=35); (3) medication management (including drug monitoring with adjustment or change in medication) (N=34); (4) measurement of BP, hypertension staging and risk of stratification, and reviewing of home BP measurements (N=13); (5) reminder system (including telephone contact, web services, home visits, or drug adherence aid) (N=12); and (6) healthcare professional education (including training program) (N=2).

The frequency of intervention was monthly or more frequently in 17 studies and less frequently than once a month in 11 studies. In the 11 remaining studies, the frequency of intervention was irregular (e.g., at every patient encounter) or not clearly specified.

Overall, study quality was moderate (Figure S1), with considerable variation in quality. Sixteen studies were considered of relatively high quality. In all studies, the participants were not blinded to the intervention.

Main Results

The outcome included systolic BP in 39 studies, involving 14 224 patients, and diastolic BP in 36 studies, involving 13 826 patients. On average, pharmacist interventions were associated with a large reduction in systolic and diastolic BP of -7.6 mm Hg (95% CI: -9.0 to -6.3 mm Hg) and

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-3.9 mm Hg (95% CI: -5.1 to -2.8 mm Hg), respectively (Figures 1A and 1B). There was a substantial heterogeneity for systolic (I²=67%) and diastolic BP (I²=83%). The 95% PI ranged from -13.9 to -1.4 mm Hg for systolic BP and from -9.9 to +2.0 mm Hg for diastolic BP.

Subgroup Analyses

Results of subgroup analyses according to selected study characteristics are summarized in Table. No difference was observed in BP reductions according to the country where the study was conducted, to the type of interventions (data not shown), to the involvement or not of a nurse in the intervention, or to the inclusion or not of patients with diabetes. The effect was slightly better if the intervention was conducted in community pharmacies. A more substantial difference in effect on BP was observed according to the type of pharmacist care, pharmacist-led care being associated with a larger effect on systolic and diastolic BP compared with collaborative care (Table, Figures 2A and 2B; systolic BP: -8.5 mm Hg [95% Cl, -10.0 to -7.0 mm Hg] versus -6.3 mm Hg [95% Cl, -8.0 to -4.5], *P*=0.046; diastolic BP: -4.6 mm Hg [95% Cl, -5.7 to -3.4 mm Hg] versus -2.8 mm Hg [95% Cl, -4.4 to -1.2 mm Hg], *P*=0.10). Furthermore, the effect on BP tended to be more important if the intervention was conducted monthly or more frequently compared with less frequently than once a month (Table, Figures 3A and 3B; systolic BP: -9.1 mm Hg [95% Cl, -9.1 to -6.7 mm Hg] versus -6.7 mm Hg [95% Cl, -9.1 to -4.4 mm Hg], *P*=0.14; diastolic BP: -4.5 mm Hg [95% Cl, -9.1 to -4.4 mm Hg], *P*=0.14; diastolic BP: -4.5 mm Hg [95% Cl, -9.1 to -4.4 mm Hg], *P*=0.14; diastolic BP: -4.5 mm Hg [95% Cl, -9.1 to -4.4 mm Hg], *P*=0.14; diastolic BP: -4.5 mm Hg [95% Cl, -9.1 to -4.4 mm Hg], *P*=0.14; diastolic BP: -4.5 mm Hg [95% Cl, -9.1 to -4.4 mm Hg], *P*=0.14; diastolic BP: -4.5 mm Hg [95% Cl, -9.1 to -4.4 mm Hg], *P*=0.14; diastolic BP: -4.5 mm Hg [95% Cl, -9.1 to -4.4 mm Hg], *P*=0.14; diastolic BP: -4.5 mm Hg [95% Cl, -9.1 to -4.4 mm Hg], *P*=0.14; diastolic BP: -4.5 mm Hg [95% Cl, -9.1 to -4.4 mm Hg], *P*=0.14; diastolic BP: -4.5 mm Hg [95% Cl, -9.1 to -4.4 mm Hg], *P*=0.14; diastolic BP: -4.5 mm Hg [95% Cl, -9.1 to -4.4 mm Hg], *P*=0.14; diastolic BP: -4.5 mm Hg]

(A)	Pharmacia	lleual						Moon difforonce	Woigh
Study	care (n)	care (n)				_		(95% CI)	(%)
Al Mazroui et al, 2009	117	117			+•	►		-4.90 (-8.43, -1.37)	3.72
Bogden et al, 1998	49	46			•	= I		-12.00 (-21.06, -2.94)	1.50
Borenstein et al, 2003	98	99			•	-		-11.00 (-18.60, -3.40)	1.90
Carter et al, 2008	101	78				-		-8.70 (-14.31, -3.09)	2.65
Carter et al, 2009	192	210			+	_		-10.90 (-14.43, -7.37)	3.72
Chan et al, 2012	51	54				•		-3.30 (-10.32, 3.72)	2.09
Chiu et al, 2008	78	76			← :			-12.80 (-17.19, -8.41)	3.25
Clifford et al, 2005	92	88			+	_		-7.00 (-13.64, -0.36)	2.23
De Castro et al, 2006	30	34			-	•		-2.00 (-7.60, 3.60)	2.66
Edelman et al, 2010	133	106			•	_		-7.30 (-15.11, 0.51)	1.83
Fornos et al, 2006	56	56						-15.00 (-21.75, -8.25)	2.18
Garçao et al, 2002	41	41		•	_			-18.36 (-29.96, -6.76)	1.04
Green et al, 2008	261	259				-		-6.00 (-9.35, -2.65)	3.82
Hennesy et al, 2006	3617	3542				•		-3.00 (-3.81, -2.19)	4.97
Hunt et al, 2008	233	230						-6.00 (-9.75, -2.25)	3.60
Jarab et al, 2012	77	79				F I I		-6.90 (-9.62, -4.18)	4.17
Kraemer et al, 2012	36	27		-	++			-5.90 (-16.26, 4.46)	1.23
Lee et al, 2006	73	62		-	•	- 1		-8.90 (-15.14, -2.66)	2.38
Magid et al, 2011	138	145			-+	H		-6.00 (-10.25, -1.75)	3.33
Magid et al, 2013	175	173		-	••• i			-12.40 (-15.80, -9.00)	3.79
Margolis et al, 2013	228	222			- + -	·		-9.70 (-14.91, -4.49)	2.84
McKenney et al, 1973	24	25		•				-20.00 (-30.47, -9.53)	1.21
McLean et al, 2008	115	112			+	H		-5.60 (-9.71, -1.49)	3.40
Mehos et al, 2000	18	18			•	-		-10.10 (-20.62, 0.42)	1.21
Morgado et al, 2001	98	99				-		-6.80 (-11.42, -2.18)	3.13
Mourao et al, 2013	50	50			•			-9.20 (-18.18, -0.22)	1.52
Okamoto et al, 2001	164	166				•		-7.80 (-11.34, -4.26)	3.71
Planas et al, 2009	25	15		•		-		-20.05 (-37.27, -2.83)	0.53
Rothman et al, 2005	99	95			•			-9.00 (-18.09, 0.09)	1.49
Santschi et al, 2008	34	34			++			-5.50 (-14.04, 3.04)	1.63
Scott et al, 2006	64	67				- L -		-5.50 (-8.69, -2.31)	3.91
Simpson et al, 2011	131	129				►		-4.90 (-10.30, 0.50)	2.75
Solomon et al, 1998	63	70			+			-6.40 (-12.46, -0.34)	2.46
Sookaneknun et al, 2004	118	117			→	►		-4.65 (-9.35, 0.05)	3.09
Taveira et al, 2010	58	51				F+ -		-5.60 (-13.10, 1.90)	1.93
Vivian et al, 2002	26	27				-		-14.40 (-24.96, -3.84)	1.20
Wang et al, 2011	29	30		-	•			-8.90 (-16.50, -1.30)	1.90
Zhao et al, 2012	129	129						-6.70 (-10.82, -2.58)	3.39
Zillich et al, 2005	64	61						-4.50 (-10.15, 1.15)	2.64
Overall (I-squared = 66.8	%, p = 0.0	00)		-				-7.64 (-8.96, -6.32)	100.00
with estimated predictive	interval				i I			. (-15.84, 0.56)	
					i	_			
			-30	-20	-10	0	10	[mmHg]	
			Favo	ors pharmad	ist care		Fa	avors usual care	

Figure 1. Forest plot of the mean difference in (A) systolic and (B) diastolic blood pressure with pharmacist care compared with usual care group. n=number of participants.

Study	Pharmacist care (n)	t Usual care (n)								Mean difference (95% CI)	Weight (%)
Al Mazroui et al, 2009	117	117								-7.80 (-9.98, -5.62)	3.43
Bogden et al, 1998	49	46		-	•	1				-11.00 (-15.43, -6.57)	2.43
Borenstein et al, 2003	98	99				! +	•			1.00 (-2.13, 4.13)	3.01
Carter et al, 2008	101	78				<u>- </u>				-5.40 (-9.06, -1.74)	2.76
Carter et al, 2009	192	210				-	-			-1.40 (-3.90, 1.10)	3.29
Chan et al, 2012	51	54			-	<u>+ • </u>	-			-2.10 (-5.46, 1.26)	2.90
Chiu et al, 2008	78	76				+				-6.70 (-9.74, -3.66)	3.05
Clifford et al, 2005	92	88			_	•				-3.00 (-6.48, 0.48)	2.85
De Castro et al, 2006	30	34			-			•		0.00 (-5.60, 5.60)	1.97
Edelman et al, 2010	133	106				•				-3.80 (-8.09, 0.49)	2.49
Fornos et al, 2006	56	56			-	•				-3.90 (-7.18, -0.62)	2.94
Garçao et al, 2002	41	41		_	•					-7.02 (-12.69, -1.35)	1.95
Green et al, 2008	261	259				-				-2.60 (-4.06, -1.14)	3.70
Hennesy et al, 2006	3617	3542				i .	•			0.00 (-0.48, 0.48)	3.92
Hunt et al, 2008	233	230								-3.00 (-5.01, -0.99)	3.50
Jarab et al, 2012	77	79		_	•	·:				-8.90 (-12.90, -4.90)	2.61
Kraemer et al, 2012	36	27				·				-1.90 (-8.58, 4.78)	1.63
_ee et al, 2006	73	62				+	_			-1.10 (-4.56, 2.36)	2.86
Magid et al, 2011	138	145								-2.30 (-5.20, 0.60)	3.11
Magid et al, 2013	175	173			-+	H				-5.70 (-7.84, -3.56)	3.45
Margolis et al, 2013	228	222								-5.10 (-8.33, -1.87)	2.96
McKenney et al, 1973	24	25			•	1 1				-11.00 (-16.77, -5.23)	1.91
Vehos et al, 2000	18	18		-	•	÷				-6.70 (-12.19, -1.21)	2.01
Vorgado et al, 2011	98	99				+++	-			-1.90 (-4.96, 1.16)	3.04
Mourao et al, 2013	50	50				+ •				-0.50 (-4.93, 3.93)	2.43
Okamoto et al, 2001	164	166			-	•				-3.60 (-5.68, -1.52)	3.47
Rothman et al, 2005	99	95								-5.00 (-10.57, 0.57)	1.98
Santschi et al, 2008	34	34					•			2.70 (-3.63, 9.03)	1.73
Simpson et al, 2011	131	129			_	•	•			-2.90 (-6.69, 0.89)	2.71
Solomon et al, 1998	63	70			_					-3.00 (-6.59, 0.59)	2.80
Sookaneknun et al, 2004	118	117								-2.48 (-5.20, 0.24)	3.19
Taveira et al, 2010	58	51			•	-i				-7.50 (-11.43, -3.57)	2.65
/ivian et al, 2002	26	27		•	_	1.00				-14.90 (-20.37, -9.43)	2.02
Nang et al, 2011	29	30			_	•				-3.60 (-6.99, -0.21)	2.89
Zhao et al, 2012	129	129			-	•				-2.90 (-5.58, -0.22)	3.21
Zillich et al, 2005	64	61			-	•				-3.20 (-6.04, -0.36)	3.14
Overall (I-squared = 83.2	%, p = 0.0	00)				ົ∽⊢				-3.94 (-5.05, -2.83)	100.0
with estimated predictive	interval									(-11.36, 3.47)	
		1	I			-					
		-30	-20		-10	0		10) [mmHa	1	

Figure 1. Continued.

-6.1 to -2.9] versus -1.9 mm Hg [95% Cl, -3.5 to -0.3 mm Hg], *P*=0.08). Nevertheless, the differences between these subgroups reached statistical significance (*P*<0.05) only for the difference in systolic BP according to the type of pharmacist care. Furthermore, the between-study heterogeneity within subgroups remained substantial in most analyses.

Publication Bias and Sensitivity Analysis

Asymmetries in the funnel plots were observed for both systolic and diastolic BP. The Egger test results were statistically significant (P<0.001 for systolic BP; P<0.001 for diastolic BP). To account for this potential publications bias and to evaluate the impact of relatively small studies, we conducted a first sensitivity analysis restricted to the 19 studies with >80 patients per arm. For these studies, the weighted mean difference between pharmacist intervention and usual care was -7.0 mm Hg (95% Cl, -8.7 to

-5.4 mm Hg; I²=72%) for systolic BP and -3.1 mm Hg (95% Cl, -4.4 to -1.8 mm Hg; I²=84%) for diastolic BP. These differences were of the same magnitude compared with the differences observed when all studies were included.

The second sensitivity analysis was restricted to the 16 studies of relatively high quality. For these studies, the weighted mean difference between pharmacist intervention and usual care was -7.3 mm Hg (95% Cl, $-8.6 \text{ to} -6.0 \text{ mm Hg}; \text{l}^2=8\%$) for systolic BP and -3.6 mm Hg (95% Cl, $-4.6 \text{ to} -2.6 \text{ mm Hg}; \text{l}^2=41\%$) for diastolic BP. These differences were of the same magnitude compared with the differences observed when all studies were included.

Discussion

By updating and combining data from our previous systematic reviews, we analyzed data from 39 RCTs including 14 224 outpatients. Findings from the current review are similar to **Table.** Subgroup Analyses for the Difference in Systolic and Diastolic Blood Pressure With Pharmacist Care Compared With Usual Care Group According to Study Characteristics

	Systolic Blood Pressure Mean Br Difference (mm Hg) D		Between Subgroup Differences	Diast Differ	olic Blood Pressure Mean ence (mm Hg)	Between Subgroup Difference
Study Characteristics	N	(95% CI)	P Value	N	(95% CI)	P Value
Country						
US/Canada	25	-7.8 (-9.6 to -6.1)	Ref	22	-4.1 (-5.5 to -2.7)	Ref
Other countries	14	-7.3 (-9.2 to -5.4)	0.83	14	-3.7 (-5.3 to -2.2)	0.76
Setting						
Outpatient clinics/GP	31	-7.5 (-8.9 to -6.1)	Ref	29	-4.0 (-5.2 to -2.7)	Ref
Community pharmacy	6	-10.5 (-15.7 to -5.3)	0.29	5	-5.0 (-7.6 to -2.3)	0.51
Other	2	-4.9 (-9.0 to -0.7)	0.39	2	-0.7 (-2.5 to 2.7)	0.22
Targeted patients		<u>^</u>	-		<u>.</u>	<u>^</u>
Without diabetes	25	-8.2 (-10.0 to -6.3)	Ref	25	-3.7 (-5.0 to -2.5)	Ref
With diabetes	14	-6.4 (-7.8 to -5.1)	0.37	11	-4.5 (-6.3 to -2.8)	0.51
Pharmacist care						
Collaborative care	16	-6.3 (-8.0 to -4.5)	Ref	14	-2.8 (-4.4 to -1.2)	Ref
Pharmacist-led care	23	-8.5 (-10.0 to -7.0)	0.046	22	-4.6 (-5.7 to -3.4)	0.10
Nurse interventions						
Without nurse	33	-7.8 (-9.3 to -6.3)	Ref	32	-4.0 (-5.1 to -2.8)	Ref
With nurse	6	-6.5 (-8.5 to -4.4)	0.65	4	-3.8 (-7.7 to 0.1)	0.92
Frequency of intervention						
Once a month or more frequently	17	-9.1 (-11.4 to -6.7)	Ref	16	-4.5 (-6.1 to -2.9)	Ref
Less than once a month	11	-6.7 (-9.1 to -4.4)	0.14	9	-1.9 (-3.5 to -0.3)	0.08
Irregular or not clearly specified	11	-7.1 (-8.6 to -5.6)	0.22	11	-4.6 (-6.1 to -3.0)	0.93

N indicates number of studies.

those of the 2 previous systematic reviews: positive evidence was found supporting pharmacist interventions for improving systolic and diastolic BP control among outpatients. The observed effect was indeed substantial compared with powerful antihypertensive drugs like loop diuretics that decrease systolic blood pressure, on average, by 8 mm Hg.⁵⁶ The interventions – conducted by the pharmacist alone or in collaboration with physician, nurse or dietitian, or physical therapist - were most often educational interventions to patients about medication, lifestyle and physical activity, feedback to physician (such as recommendation regarding medications changes or problems to medication adherence), and medication management (such as drug monitoring with adjustment or change in medication). A substantial heterogeneity in the effect estimate was observed for both systolic and diastolic BP. Although our subgroup analyses could not explain this heterogeneity, interventions led by the pharmacist tended to be more effective. Furthermore, interventions conducted at least monthly might be more effective, albeit the difference did not reach statistical significance.

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Finally, wide prediction intervals were observed, suggesting differential effects of pharmacist interventions on BP in individual studies, from a very large effect to modest or no effect.

Findings from the current systematic review and metaanalysis are consistent with previous reviews showing that pharmacist-led care or involved in a team-based care with physician or nurse improves patient health outcomes^{6,57} including BP control.^{5,9,10} In most RCTs included in this updated and combined review, complex pharmacist interventions with multiple components (i.e., co-interventions) were evaluated, making it difficult to identify precisely which component of the intervention was more potent to control BP. Despite the large number of included studies, our analyses did not clearly identify which (aspect of the) intervention is the most efficient. We have shown in a previous review that the number of interventions was not associated with a better BP control.9 In this current combined review, we observed a slightly larger effect on BP with pharmacist-led care compared with pharmacist interventions implemented in collaboration

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Study	care (n) care (n)					(95% CI)	(%)
Pharmacist-led care								
Al Mazroui et al, 2009	117	117					-4.90 (-8.43, -1.37)	3.72
Chan et al, 2012	51	54				•	-3.30 (-10.32, 3.72)	2.09
Chiu et al, 2008	78	76		- E	•		-12.80 (-17.19, -8.41)	3.25
Fornos et al, 2006	56	56			—		-15.00 (-21.75, -8.25)	2.18
Garçao et al, 2002	41	41		•			-18.36 (-29.96, -6.76)	1.04
Green et al, 2008	261	259					-6.00 (-9.35, -2.65)	3.82
Jarab et al, 2012	77	79				· · · ·	-6.90 (-9.62, -4.18)	4.17
Kraemer et al, 2012	36	27		-	. •	<u> </u>	-5.90 (-16.26, 4.46)	1.23
Lee et al, 2006	73	62			•	-	-8.90 (-15.14, -2.66)	2.38
Magid et al, 2013	175	173		-			-12.40 (-15.80, -9.00)	3.79
Margolis et al, 2013	228	222					-9.70 (-14.91, -4.49)	2.84
McKenney et al, 1973	24	25	-	•			-20.00 (-30.47, -9.53)	1.21
Mehos et al, 2000	18	18			•		-10.10 (-20.62, 0.42)	1.21
Morgado et al, 2001	98	99			•	-	-6.80 (-11.42, -2.18)	3.13
Mourao et al. 2013	50	50			•		-9.20 (-18.18, -0.22)	1.52
Okamoto et al. 2001	164	166					-7.80 (-11.34, -4.26)	3.71
Planas et al. 2009	25	15		•	-	-	-20.05 (-37.27, -2.83)	0.53
Rothman et al. 2005	99	95			•	_	-9.00 (-18.09, 0.09)	1.49
Sookaneknun et al. 2004	118	117			-+-+		-4.65 (-9.35, 0.05)	3.09
Vivian et al. 2002	26	27			•		-14.40 (-24.96, -3.84)	1.20
Wang et al. 2011	29	30			•		-8.90 (-16.50, -1.30)	1.90
Zhao et al. 2012	129	129				- 1	-6.70 (-10.82, -2.58)	3.39
Zillich et al. 2005	64	61					-4.50 (-10.15, 1.15)	2.64
Subtotal (I-squared = 42.	4%. p =	0.017)		-	\rightarrow		-8.50 (-10.03, -6.97)	55.54
with estimated predictive	interval	,					. (-16.54, -0.46)	
Collaborative care					i i i			
Bogden et al, 1998	49	46				-	-12.00 (-21.06, -2.94)	1.50
Borenstein et al, 2003	98	99			•	-	-11.00 (-18.60, -3.40)	1.90
Carter et al, 2008	101	78			•	-	-8.70 (-14.31, -3.09)	2.65
Carter et al, 2009	192	210					-10.90 (-14.43, -7.37)	3.72
Clifford et al, 2005	92	88					-7.00 (-13.64, -0.36)	2.23
De Castro et al, 2006	30	34				•	-2.00 (-7.60, 3.60)	2.66
Edelman et al, 2010	133	106			•		-7.30 (-15.11, 0.51)	1.83
Hennesy et al, 2006	3617	3542			- 1	•	-3.00 (-3.81, -2.19)	4.97
Hunt et al, 2008	233	230			•		-6.00 (-9.75, -2.25)	3.60
Magid et al, 2011	138	145			-		-6.00 (-10.25, -1.75)	3.33
McLean et al, 2008	115	112				H .	-5.60 (-9.71, -1.49)	3.40
Santschi et al, 2008	34	34			+		-5.50 (-14.04, 3.04)	1.63
Scott et al, 2006	64	67				- 1	-5.50 (-8.69, -2.31)	3.91
Simpson et al, 2011	131	129				H	-4.90 (-10.30, 0.50)	2.75
Solomon et al, 1998	63	70				_	-6.40 (-12.46, -0.34)	2.46
Taveira et al, 2010	58	51			•		-5.60 (-13.10, 1.90)	1.93
Subtotal (I-squared = 58.	1%, p =	0.002)			\rightarrow		-6.25 (-7.96, -4.53)	44.46
with estimated predictive	interval				- I -		. (-15.30, 2.81)	
Overall (I-squared = 66.8	8%, p =	0.000)			Ŷ		-7.64 (-8.96, -6.32)	100.00
with estimated predictive	interval						. (-15.84, 0.56)	
					i			
			1					_
			-30	-20	-10	0	10 [mmHg]	
			Fav	ors nharm	acist care		Favors usual care	
			iav					

Figure 2. Forest plot of the mean difference in (A) systolic and (B) diastolic blood pressure with pharmacist care compared with usual care group according to type of care (pharmacist-led care vs collaborative care). n=number of participants.

with other health professionals. Nevertheless, the difference was modest.

Our findings that the effect on BP was slightly greater if the intervention was led by the pharmacist differ to some extend from a recent umbrella review of systematic reviews evaluating the effectiveness of community pharmacist intervention.⁵⁸ Mossialos et al⁵⁸ reported mixed or inconclusive evidence in support of expanding the role of community pharmacist in healthcare delivery. One reason for the difference may be that the review by Mossialos et al⁵⁸ was focused specifically on community pharmacists. Our review

evaluated not only studies with community pharmacist interventions but also clinical pharmacist interventions working, for example in outpatient clinic or medical groups. Nevertheless, we agree with Mossialos et al⁵⁸ that expanding the role of community pharmacist needs significant changes in healthcare systems, further training of pharmacists and involves policy development.

Furthermore, based on a thorough review, the team-based hypertension care has been recently recommended by the US Preventive Services Task Forces as an effective strategy to improve BP, with nonphysician healthcare

(A)									
Study P	harmac care (n	ist Usual) care (n)					Mean difference (95% CI)	Weight (%)
Pharmacist-led care					!				
Al Mazroui et al, 2009	117	117			+●	- 1		-4.90 (-8.43, -1.37)	3.72
Chan et al, 2012	51	54			<u>— — — — — — — — — — — — — — — — — — — </u>	•		-3.30 (-10.32, 3.72)	2.09
Chiu et al, 2008	78	76				_		-12.80 (-17.19, -8.41)	3.25
Fornos et al, 2006	56	56			•i			-15.00 (-21.75, -8.25)	2.18
Garcao et al, 2002	41	41			_			-18.36 (-29.96, -6.76)	1.04
Green et al, 2008	261	259				H I		-6.00 (-9.35, -2.65)	3.82
Jarab et al, 2012	77	79						-6.90 (-9.62, -4.18)	4.17
Kraemer et al, 2012	36	27		-	+			-5.90 (-16.26, 4.46)	1.23
Lee et al, 2006	73	62			-	-		-8.90 (-15.14, -2.66)	2.38
Magid et al, 2013	175	173						-12.40 (-15.80, -9.00)	3.79
Margolis et al, 2013	228	222		_				-9.70 (-14.91, -4.49)	2.84
McKenney et al, 1973	24	25		•				-20.00 (-30.47, -9.53)	1.21
Mehos et al, 2000	18	18			• <u>'</u>	-		-10.10 (-20.62, 0.42)	1.21
Morgado et al, 2001	98	99				-		-6.80 (-11.42, -2.18)	3.13
Mourao et al, 2013	50	50		_	•			-9.20 (-18.18, -0.22)	1.52
Okamoto et al, 2001	164	166						-7.80 (-11.34, -4.26)	3.71
Planas et al, 2009	25	15		•		-		-20.05 (-37.27, -2.83)	0.53
Rothman et al, 2005	99	95		_	•			-9.00 (-18.09, 0.09)	1.49
Sookaneknun et al, 2004	118	117			4			-4.65 (-9.35, 0.05)	3.09
Vivian et al, 2002	26	27			• •			-14.40 (-24.96, -3.84)	1.20
Wang et al, 2011	29	30		-	•	-		-8.90 (-16.50, -1.30)	1.90
Zhao et al, 2012	129	129			-	- 1		-6.70 (-10.82, -2.58)	3.39
Zillich et al, 2005	64	61						-4.50 (-10.15, 1.15)	2.64
Subtotal (I-squared = 42.	4%, p =	0.017)		-		_		-8.50 (-10.03, -6.97)	55.54
with estimated predictive	interval				i i			. (-16.54, -0.46)	
Collaborative care									
Bogden et al. 1998	49	46			•	-		-12.00 (-21.06, -2.94)	1.50
Borenstein et al. 2003	98	99			•	-		-11.00 (-18.60, -3.40)	1.90
Carter et al 2008	101	78				-		-8 70 (-14 31 -3 09)	2 65
Carter et al. 2009	192	210						-10.90 (-14.43, -7.37)	3.72
Clifford et al. 2005	92	88						-7.00 (-13.64, -0.36)	2.23
De Castro et al. 2006	30	34				•		-2.00 (-7.60, 3.60)	2.66
Edelman et al. 2010	133	106				-		-7.30 (-15.11, 0.51)	1.83
Hennesv et al. 2006	3617	3542				•		-3.00 (-3.81, -2.19)	4.97
Hunt et al. 2008	233	230				- 1		-6.00 (-9.75, -2.25)	3.60
Magid et al, 2011	138	145				- 1		-6.00 (-10.25, -1.75)	3.33
McLean et al, 2008	115	112				-1		-5.60 (-9.71, -1.49)	3.40
Santschi et al, 2008	34	34						-5.50 (-14.04, 3.04)	1.63
Scott et al, 2006	64	67				- L -		-5.50 (-8.69, -2.31)	3.91
Simpson et al, 2011	131	129						-4.90 (-10.30, 0.50)	2.75
Solomon et al, 1998	63	70			-			-6.40 (-12.46, -0.34)	2.46
Taveira et al, 2010	58	51				H-		-5.60 (-13.10, 1.90)	1.93
Subtotal (I-squared = 58.	1%, p =	0.002)			\rightarrow	<u> </u>		-6.25 (-7.96, -4.53)	44.46
with estimated predictive	interval	,			1			. (-15.30, 2.81)	
Overall (Leauared - 66 9	% n =	0 000)			<u> </u>			-7 64 (-8 96 -6 32)	100.00
with estimated predictive	interval	0.000)			Ť			(-15.84 0.56)	100.00
mui esumateu preuictive	niicivdi				1			. (-13.04, 0.30)	
					'		1		
			-30	-20	-10	0	10 [mr	nHg]	
			Fa	vors pharm	acist care		Favo	rs usual care	
				•					

Figure 2. Contiuned.

professionals – involving pharmacist, nurses, dietitian – working in collaboration with physician and sharing responsibilities of hypertension care.⁷ Comparative effectiveness studies are therefore needed to determine the most efficient interventions and long-term effect.

Potentially important components of effective pharmacist interventions such as adequate pharmacist training (e.g., measurement of BP or management of hypertension) and remuneration for services (e.g., reviews of medications, resolution of drug related-problems, patient counseling to improve adherence) should be considered to better understand the heterogeneity of the results of pharmacist interventions reported here. Another possible factor that might explain the reported difference in the reviewed papers is the various types of healthcare delivery system. Indeed, pharmacist involvement is typical and well accepted in the care, for example in the United States or in Canada, compared with European countries, and that could influence the effect size of pharmacist interventions. Obstacles to the implementation of intervention should be also considered. It would be especially relevant for policy recommendations to have such information. However, and this is one limitation of the current

(A)								
Study	Pharmacist U care (n) ca	sual e (n)					Mean difference (95% CI)	Weight (%)
Once a month or more	frequently							
Borenstein et al, 2003	98 99						-11.00 (-18.60, -3.40)	1.90
Chiu et al, 2008	78 76			•			-12.80 (-17.19, -8.41)	3.25
De Castro et al, 2006	30 34				•		-2.00 (-7.60, 3.60)	2.66
Fornos et al, 2006	56 56						-15.00 (-21.75, -8.25)	2.18
Garçao et al, 2002	41 41						-18.36 (-29.96, -6.76)	1.04
Green et al, 2008	261 259						-6.00 (-9.35, -2.65)	3.82
Makannay at al. 1072	228 222						-9.70 (-14.91, -4.49)	2.84
Mohoo et al. 2000	24 23 10 10						-20.00 (-30.47, -9.53)	1.21
Mourao et al. 2000	50 50				_		-9.20 (-18.18 -0.22)	1.21
Planas et al. 2009	25 15		•		.		-20.05 (-37.27 -2.83)	0.53
Rothman et al. 2005	99 95			•	_		-9.00 (-18.09, 0.09)	1.49
Solomon et al. 1998	63 70						-6.40 (-12.46, -0.34)	2.46
Sookaneknun et al, 2004	118 117				н		-4.65 (-9.35, 0.05)	3.09
Taveira et al, 2010	58 51				-		-5.60 (-13.10, 1.90)	1.93
Vivian et al, 2002	26 27						-14.40 (-24.96, -3.84)	1.20
Zillich et al, 2005	64 61			+	H		-4.50 (-10.15, 1.15)	2.64
Subtotal (I-squared = 48	.6%, p = 0.01	3)		\rightarrow			-9.05 (-11.35, -6.74)	34.96
with estimated predictive	interval						. (-21.17, 3.08)	
Less than once a mont	h							
Carter et al, 2008	101 78						-8.70 (-14.31, -3.09)	2.65
Carter et al, 2009	192 210						-10.90 (-14.43, -7.37)	3.72
Clifford et al, 2005	92 88		-				-7.00 (-13.64, -0.36)	2.23
Edelman et al. 2010	133 100	`		The second se	•		-7.30 (-15.11, 0.51)	1.83
Kraomor et al. 2000	36 27	2	_				-3.00 (-3.01, -2.19)	4.97
Lee et al 2006	73 62		-		-		-8.90 (-15.14 -2.66)	2.38
Mclean et al 2008	115 112				-		-5 60 (-9 71 -1 49)	3 40
Santschi et al. 2008	34 34						-5.50 (-14.04, 3.04)	1.63
Scott et al. 2006	64 67				-		-5.50 (-8.69, -2.31)	3.91
Wang et al, 2011	29 30		_	•			-8.90 (-16.50, -1.30)	1.90
Subtotal (I-squared = 66	.9%, p = 0.00	1)		\rightarrow	_		-6.74 (-9.07, -4.41)	29.86
with estimated predictive	interval			i i			. (-19.30, 5.81)	
Irregularly or not speci	fied							
Al Mazroui et al, 2009	117 117						-4.90 (-8.43, -1.37)	3.72
Bogden et al, 1998	49 46						-12.00 (-21.06, -2.94)	1.50
Chan et al, 2012	51 54						-3.30 (-10.32, 3.72)	2.09
Hunt et al, 2008	233 230						-6.00 (-9.75, -2.25)	3.60
Jarab et al, 2012	// /9						-6.90 (-9.62, -4.18)	4.17
Magid et al. 2012	138 145		_				-0.00 (-10.25, -1.75)	3.33
Morgado et al. 2013	08 00			-	-		-12.40 (-15.60, -9.00)	3.19
Okamoto et al. 2001	90 99 164 166						-0.00 (-11.42, -2.16)	3.13
Simpson et al. 2001	131 129				4		-4 90 (-10 30 0 50)	2 75
Zhao et al 2012	129 129						-6 70 (-10 82 -2 58)	3 39
Subtotal (I-squared = 31	.0%, p = 0.15	2)		$-\diamond$	-		-7.09 (-8.60, -5.58)	35.19
with estimated predictive	interval	,		- I			. (-14.79, 0.61)	
Overall (I-squared = 66.)	8%, p = 0.000)	_				-7.64 (-8.966.32)	100.00
with estimated predictive	interval			Ï			. (-15.84, 0.56)	
		I 20	I 20	10	I	10 [mm	Hal	
		-30	-20	-10	U	10 [1111		
		Fa	avors pharm	acist care		Favors	s usual care	

Figure 3. Forest plot of the mean difference in (A) systolic and (B) diastolic blood pressure with pharmacist care compared with usual care group according to frequency of intervention (once a month or more frequently, less than once a month, irregularly/not specified). n=number of participants.

review, information about these factors was not available or only partly reported in the studies identified.

To better express uncertainties in the effect estimates of pharmacist intervention, we reported prediction intervals.^{15,16} When there is substantial heterogeneity, the interpretation of the mean effect and the related confidence interval may be problematic, especially when the statistical heterogeneity is due to true clinical heterogeneity.¹⁵ Prediction interval reveals the distribution of effect estimates for individual studies, and our results suggest that there could be large differences in

effect estimates, with some intervention having modest or no effect on BP. Prediction interval gives the range of effects that would be expected from a new study with similar characteristics to the current studies. Reporting prediction intervals can be enlightening especially when complex healthcare interventions, with expected true clinical heterogeneity in their effects, are evaluated.

Some limitations should be considered in interpreting the results of this study. Asymmetries in funnel plots were clearly observed for systolic and diastolic BP, and this may reflect a

B)							
Study F	harmacis care (n)	t Usual care (n)				Mean difference (95% CI)	Weight (%)
Once a month or more	frequent	tly					
Borenstein et al, 2003	98	99				1.00 (-2.13, 4.13)	3.01
Chiu et al, 2008	78	76				-6.70 (-9.74, -3.66)	3.05
De Castro et al, 2006	30	34		_	•	0.00 (-5.60, 5.60)	1.97
Fornos et al, 2006	56	56			-	-3.90 (-7.18, -0.62)	2.94
Garçao et al, 2002	41	41			_	-7.02 (-12.69, -1.35)	1.95
Green et al, 2008	261	259			•	-2.60 (-4.06, -1.14)	3.70
Margolis et al, 2013	228	222		-+	_	-5.10 (-8.33, -1.87)	2.96
McKenney et al, 1973	24	25		•		-11.00 (-16.77, -5.23)	1.91
Mehos et al, 2000	18	18				-6.70 (-12.19, -1.21)	2.01
Mourao et al, 2013	50	50			•	-0.50 (-4.93, 3.93)	2.43
Rothman et al, 2005	99	95			-	-5.00 (-10.57, 0.57)	1.98
Solomon et al, 1998	63	70				-3.00 (-6.59, 0.59)	2.80
Sookaneknun et al, 200	4 118	117		7	•	-2.48 (-5.20, 0.24)	3.19
Taveira et al, 2010	58	51				-7.50 (-11.43, -3.57)	2.65
Vivian et al, 2002	26	27				-14.90 (-20.37, -9.43)	2.02
Zillich et al, 2005	64	61			•	-3.20 (-6.04, -0.36)	3.14
Subtotal (I-squared = 6	9.1%, p =	0.000)		-		-4.49 (-6.10, -2.88)	41.71
with estimated predictiv	e interval					. (-13.50, 4.52)	
Less than once a mon	th						
Carter et al, 2008	101	78				-5.40 (-9.06, -1.74)	2.76
Carter et al, 2009	192	210				-1.40 (-3.90, 1.10)	3.29
Clifford et al, 2005	92	88				-3.00 (-6.48, 0.48)	2.85
Edelman et al, 2010	133	106				-3.80 (-8.09, 0.49)	2.49
Hennesy et al, 2006	3617	3542				0.00 (-0.48, 0.48)	3.92
Kraemer et al, 2012	36	27				-1.90 (-8.58, 4.78)	1.63
Lee et al, 2006	73	62				-1.10 (-4.56, 2.36)	2.80
Santschi et al, 2006	34	34				2.70 (-3.63, 9.03)	1.73
Subtotal (Laguarad - 5	29 0.0% n -	30				-3.00 (-0.99, -0.21)	2.09
with estimated predictive	e interval	0.012)			~	. (-10.50, 6.68)	24.42
Irregularly or not spec	ified						
Al Mazroui et al, 2009	117	117				-7.80 (-9.98, -5.62)	3.43
Bogden et al, 1998	49	46	-	•		-11.00 (-15.43, -6.57)	2.43
Chan et al, 2012	51	54		_	•	-2.10 (-5.46, 1.26)	2.90
Hunt et al, 2008	233	230			•	-3.00 (-5.01, -0.99)	3.50
Jarab et al, 2012	77	79		•	_	-8.90 (-12.90, -4.90)	2.61
Magid et al, 2011	138	145			•	-2.30 (-5.20, 0.60)	3.11
Magid et al, 2013	175	173				-5.70 (-7.84, -3.56)	3.45
Morgado et al, 2011	98	99		-	•	-1.90 (-4.96, 1.16)	3.04
Okamoto et al, 2001	164	166				-3.60 (-5.68, -1.52)	3.47
Simpson et al, 2011	131	129			•	-2.90 (-6.69, 0.89)	2.71
Zhao et al, 2012	129	129		_	•	-2.90 (-5.58, -0.22)	3.21
Subtotal (I-squared = 7	1.2%, p =	0.000)		\rightarrow		-4.55 (-6.11, -2.99)	33.86
with estimated predictiv	e interval					(-13.29, 4.20)	
Overall (I-squared = 83	.2%, p = 0	0.000)		\rightarrow	≻——	-3.94 (-5.05, -2.83)	100.00
with estimated predictive	e interval					(-11.36, 3.47)	
		I		1		1	
		-30	-20	-10	0	10 [mmHg]	
			Fourier	ormonict and		Fovers your error	
			Favors ph	annacist care		ravors usual care	

Figure 3. Contiuned.

publication bias.⁵⁹ Indeed, such a bias is possible because we did not search systematically for unpublished studies and because small studies showing no or weak effect of pharmacist intervention may not have been published. Nevertheless, our sensitivity analyses did not suggest any major impact of small studies on the effect estimates. Furthermore, the asymmetry may also be due to other factors, such as the heterogeneity of the effect of interventions.⁵⁹ The difference reached statistical significance only for the difference in the effect on systolic BP between collaborative care and pharmacist-led care. However, statistical tests with meta-regression in meta-analyses are known to have low power to show statistically significant differences between subgroups. The absence of statistically

significant differences does not exclude the possibility of a true and clinically relevant difference in effect size. Therefore, we considered with precaution that the effect size could differ according to the frequency of interventions and the type of care despite the absence of statistically significant differences in most comparisons.

Another limitation is that the quality of studies was moderate on average. A reassuring observation is the consistency of the effect on BP when the analysis was restricted to high-quality studies: no difference was observed compared to the effect estimates when all studies were included. The relatively short duration of follow-up (median of 6 months) is also a limitation. It is indeed possible that the effect of pharmacist interventions could be reduced over a longer duration of follow-up, as reported in a recent RCT evaluating the effect of home BP telemonitoring and pharmacist management on BP control.⁴³ Finally, a cost-effectiveness analysis of these interventions is lacking. This would help determine which intervention should be preferably implemented.

In addition to contributing to the current state of literature supporting the beneficial effects of pharmacists interventions in BP control, our findings underline that the impact of long-term (e.g., >1 year) pharmacist interventions on BP control is lacking. Moreover, most of the included studies evaluating the impact of pharmacist care on BP among outpatients were conducted in the United States and Canada, reflecting an enhanced and more successfully implemented role of pharmacist in these healthcare systems.

Conclusions

Our meta-analysis shows evidence that pharmacist interventions improve BP control in outpatients compared with usual care. Given the large heterogeneity between studies, it is difficult to predict which effect size would be expected when implementing such interventions in practice. Comparative effectiveness studies, with a long duration of follow-up, are still needed to determine the most efficient, implementable, cost-effective, and least time-consuming intervention for improving BP control in various healthcare systems.

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Disclosures

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