Running head: THE QUESTIONNAIRE BIG SIX IN 26 NATIONS

The Questionnaire Big Six (QB6) in 26 Nations: Developing Cross-Culturally

Applicable Big Six, Big Five, and Big Two Inventories

Amber Gayle Thalmayer and Gerard Saucier

University of Oregon

### Author Note

Amber Gayle Thalmayer and Gerard Saucier, Department of Psychology, University of Oregon.

Amber Gayle Thalmayer is now at OptumHealth.

Correspondence concerning this article should be addressed to Amber Gayle Thalmayer,

OptumHealth, P.O. Box 9472 OR 995-1000, Minneapolis, MN, 55440-9472. Email:

ambergayle@gmail.com

#### Abstract

The Big Five is a useful model of attributes now commonly used in cross-cultural research, but without the support of strong measurement invariance (MI) evidence. The Big Six has been proposed as a cross-culturally informed update, and the broader Big Two (Social Self-Regulation and Dynamism) draws on even more cross-cultural evidence. However, neither has been rigorously tested for cross-cultural MI. Here a Big Six inventory (36QB6) and measures of the Big Five and Big Two derived from it, were tested and refined for cross-cultural usability in samples from 26 nations, divided into three subsets. CFA of the models in the first subset of nations demonstrated fit as strong in translation as typical personality measures achieve in their nation of origin (although poor per standard benchmarks). Items that performed inconsistently across cultures were removed, and alternates considered in a second subset of nations. Fit and invariance were improved for refined 30-item QB6 (30QB6), 25-item Big Five (25QB5), and 14item Big Two (14QB2) measures in the third subset of nations. For all models, decrease in CFI between MI levels was larger than .01, indicating lack of support for higher levels. Configural and factorial invariance were relatively stronger, compared to scalar and full. Keywords: Personality Measures, Cross Cultural Psychology, Personality Traits

The Questionnaire Big Six (QB6) in 26 Nations: Developing Cross-Culturally

Applicable Big Six, Big Five and Big Two Inventories

While the scientists developing models of personality over the last century have generally come from a restricted range of nations in North America and Europe, psychology is the study of the human mind and behavior in general. It is certainly the goal of most psychologists to understand human personality in ways that transcend the immediate place and time of the researcher.

Standard practice in the field has been to thoroughly refine and validate a proposed inventory in the initial language of development. Only after an inventory is well established, is it translated into other languages, and its cross-cultural applicability assessed. With this project, we take a more 'culture-fair' approach. Data from diverse nations is drawn on to refine three personality measures (the QB6 and Big Five and Big Two measures derived from it). Here an inventory platform developed in English (but based on evidence from lexical work in many languages) is tested for cross-cultural applicability at a relatively early stage of development. This allows evidence from diverse cultures to play a role in inventory refinement, and should lead to inventories that are more culturally de-centered.

#### The Big Five

The use of lexical studies in the 1970s, based on the rationale that the most important distinctions between people will be encoded in the natural languages (Goldberg, 1981), allowed personality researchers to move away from expert judgment in selection of variables, and to base studies of personality structure on objective patterns in personality lexicons. The procedure of lexical studies is easily transferable to diverse languages. In any new setting, four basic steps can be followed: 1) extract all personality relevant terms from a dictionary; 2) reduce to a tractable

number; 3) administer in inventory-form to participants; and 4) use factor analysis to determine which terms hang together and best distinguish between individuals in the population. A standard methodology (and the linking of results by the use of marker sets) has allowed for comparison of structural models of personality attributes across languages.

The Big Five (Extraversion, Emotional Stability vs. Neuroticism, Conscientiousness, Agreeableness, and Intellect/Openness) rose to prominence when factor-analytic studies conducted with temperament and personality scales and lexical studies in English, German (Ostendorf, 1990), and Dutch (De Raad, Henriks, & Hofstee, 1992) converged on this model (Goldberg, 1990; Saucier & Goldberg, 1996). A degree of consensus has been highly generative for the field of personality psychology, and many meaningful relations between life outcomes and scores on the five factors have been established (e.g. Ozer & Benet-Martinez, 2006).

Big Five inventories have been used regularly in cross-cultural research, but have generally not been subjected to full measurement invariance analyses (Church et al, 2011). Initial investigations, including PCA, have suggested an initial level of configural invariance for the NEO-PI-R across cultures (McCrae, 1997; Poortinga, Van De Vijver, & Van Helmert, 2002). And Nye, Roberts, Saucier, & Zhou (2008), comparing one Big Five scale at a time across three cultural groups (using the mini Markers, Saucier 1994), found configural but not factorial or scalar invariance. More rigorous measurement invariance analyses of the NEO-PI-R in three nations indicated considerable differential item functioning for nearly half the items (Church et al, 2011).

#### The Big Six

Lexical studies have since been completed in languages increasingly culturally and linguistically distant from the original trio of Germanic languages. The accumulating evidence

suggests that updates to the Big Five model could make it more cross-culturally informed (Ashton et al, 2004; Saucier 2009). Studies in Italian (De Raad, DiBlas, & Perugini, 1997), Hungarian (Szirmak & De Raad, 1994), Greek (Saucier, Georgiades, Tsaousis, & Goldberg, 2005), and Chinese (Zhou, Saucier, Gao, & Liu, 2009) have not found the Big Five in the five factor solutions. The addition of a sixth factor, including content related to personal integrity versus taking advantage of others, makes the model better match empirical results from a larger group of lexical studies. A more cross-cultural model, drawn from a larger and more diverse population base, is more likely to replicate over time and across additional languages and cultures. Even in North America, the Big Six has been shown to have some theoretical and predictive advantages over the Big Five, and to contribute additional interpretive power (Ashton & Lee, 2007; Saucier, 2009; Thalmayer, Saucier, & Eigenhuis, 2011). Luckily, the Big Six model is highly isomorphic to the Big Five, allowing for straightforward integration of previous research results with this updated model.

The Questionnaire Big Six (QB6) scales, including Conscientiousness, Honesty/ Propriety, Agreeableness (Kindness & Even Temper), Resiliency versus Internalizing Negative Emotionality, Extraversion (Gregariousness and Positive Emotionality), and Originality/Talent, are highly comparable to Big Five dimensions, and to Ashton and Lee's (2007) HEXACO sixfactor inventory (Saucier, 2009). The QB6 scales are complementary to the HEXACO inventories in being shorter. Advantages of the QB6 include better elucidation of the "externalizing" domain, compared to the Big Five, because Agreeableness and Honesty/Propriety distinguish between reactive versus predatory aggression, respectively, at their low ends. Internalizing affect (depression, anxiety, tendencies toward panic and phobias) is also better represented than in Big Five measures of similar length. And Originality/Talent encompasses perceived talents, abilities, and intellectual interests, including "positive valence" content typically found in broader variable selection studies, but excluded in most inventories (Thalmayer et al, 2011).

#### The Big Two

There is evidence that one-and two-factor models of personality structure may be even more ubiquitous (Saucier, Thalmayer, & Bel-Bahar, 2014; Saucier, Thalmayer et al, 2014). Saucier, Thalmayer et al (2014) provide specific highly recurrent terms for a two dimensional model (Social Self-Regulation [S] and Dynamism [D]) drawn from nine diverse lexical studies. This 'Big Two' was not derived as higher-order factors from Big Five (or Six) scales, but from the first two factors derived when hundreds of natural-language descriptors are analyzed. One factor (D) appears to relate to the relative proportion of approach versus avoidant tendencies in the personality, whereas the other (S) relates to the internalization of social and cultural norms. This model of personality attributes is more 'culturally de-centered' – based on data from diverse populations around the world – thus it minimizes bias for or against one kind of human culture or population.

A two-factor model is the highest of three levels of structure commonly used by contemporary personality psychologists; higher-order factors of the Big Five (Digman, 1997; De Young, 2006) are similarly at this level, and are comparable to the Big Two. This simple model can be differentiated into the useful mid-level, with five or six factors. Even more predictive power is available at the facet level, where each of the five or six factors are differentiated into subcomponents (John & Srivastava, 1999), as in the NEO-PI-R (Costa & McCrae, 1992) or the HEXACO (Lee & Ashton, 2004). Psychologists will naturally choose more differentiated models where possible, to facilitate prediction. But two factors make for a parsimonious model with potential advantages for theory. They replicate reliably across diverse cultures and languages and across diverse variable selection strategies and procedures (Saucier, Thalmayer et al, 2014), and thus offer a firmer foundation for the cross-cultural study of personality. An inventory developed using this culturally de-centered model is more appropriately, and should be more easily, translated into new languages, facilitating cross-cultural research and measurement invariance.

### Cross-Cultural Measurement Invariance and Goals for the Current Study

The current study assesses cross-cultural consistency in Big Six self-report personality data from 26 nations collected as part of the Survey of World Views, a large omnibus survey of constructs relevant to cross-cultural psychology (Saucier, Kenner, et al., 2014). The 36QB6 is tested for measurement invariance (MI) and refined for cross-cultural applicability. Because the Big Five is the closest the field of personality has to a "consensual model," and because it is now often measured across cultures, a workable Big Five inventory is also constructed from items in the dataset, tested, and refined. Because the Big Two has theoretical and cross-cultural advantages, but no measure of it currently exists, a 20-item Big Two inventory developed from QB6 items is also tested and refined. Developing and validating the three measures from QB6 items can allow researchers, regardless of preferred model, to make use of the translations into 31 languages of the 40 personality items used in the Survey of World Views (all items and translations are freely available at http://psychometriglossia.uoregon.edu/).

Invariance testing allows us to determine the extent to which items are used in similar ways by different groups, and the extent to which the same patterns of correlations between items emerge. To the extent that MI can be established for the QB6 and/or it's QB5 and QB2 off-shoots, we can have more confidence that these models of personality and these specific inventories are cross-culturally appropriate. Where established, comparison of correlations

between attribute dimensions and other constructs and life outcomes of interest can confidently be made across cultural groups.

Invariance testing typically proceeds in four stages. At the most basic level, configural MI tests whether the same factors apply across groups; there are no parameter equality constraints. If established, configural MI indicates that individuals across nations use the same number of latent variables to reflect differences in scores on the items, providing a reference model for more constrained models. We can then proceed to test factorial or metric equivalence -- whether the same factor structure (number of latent variables and interrelationships to one another and indicator items) hold across the groups (Vandenberg & Lance, 2000), and thus whether same items can be used to assess the constructs across groups. Factorial invariance involves a constraint of equality-across-groups for factor loadings. If established, this suggests that items are used in a similar way across groups with respect to factor structure, and that latent variables have well-matched content across groups; in this case it is reasonable to examine the relationships of these latent variables to other constructs of interest across groups. A lack of MI at this stage means there is content in the latent constructs that varies from group to group, and suggests that items are perceived and interpreted differently, or that attributes covary with one another inconsistently across contexts. Factorial invariance is necessary though not sufficient for comparing scores across groups.

The level of scalar equivalence tests whether patterns of scores and weight parameters (factor loadings) match across groups, such that relative differences can be compared. It sets a constraint that intercepts be equal across groups, so that any cross-cultural differences cannot be attributed simply to differential functioning the single indicators in various groups. Finally, full equivalence involves constraints on the residual variances and tests whether scales measure

latent traits with similar reliability across groups. Full (or strict) invariance means that one can directly compare scores at face value across groups, interpreting differences as applied to latent constructs.

Models will initially be tested in each country using single-group confirmatory factor analysis (CFA). There is reason to anticipate levels of fit that do not achieve standard benchmarks per Hu and Bentler (1999). For multi-factor inventories like the QB6, with measurement at the item level, such benchmarks are rarely, if ever, achieved (Marsh, Hau & Wen, 2004). This might be due to a variety of factors: accumulation of item-level error, order and method effects, similarities in wording and life-domains referenced (e.g., Poortinga et al. 2002). It can also be argued that personality itself lacks the 'local independence' or simple structure that fit indices reward (Cramer et al, 2012). In the QB6, items that are direct opposites of one another have generally been avoided, and the largest possible range of domain content has been included in each short scale. Such an approach is intended to maximize predictive validity, not internal consistency.

Hopwood and Donnellan (2010) demonstrate that multidimensional personality inventories, many of which were developed by exploratory factor analysis, routinely fail to achieve adequate fit per standard benchmarks. Their CFAs of eight inventories (all constructed in North America) found that none achieved adequate fit in a North American community sample, despite established predictive validity. (Two had inadmissible results, remaining six: TLI .52-.70, CFI .61-.79, RMSEA .09-.13.) If it is difficult to achieve good fit of multidimensional models in one population, it will be even more challenging to find it in models tested across diverse populations. Thus, comparisons will be made to the fit indices reported by Hopwood and Donnellan (2010) as reasonably high, 'domain specific' benchmarks. The current study initially assesses the cross-cultural usability of the QB6. The Big Six model was developed using the results of culturally diverse lexical studies (Saucier, 2009), and thus it is expected to have a better chance of measurement invariance across cultures than many other personality models. However, the measures of the Big Six were developed using North American data as a baseline, so the QB6 is likely to fit best in the North American populations that had the most impact on its development. It is likely that it will fit less well as samples more culturally distant from this place of origin are tested.

A Big Five and a Big Two model will also be tested. The Big Five is a simpler structure than the Big Six, but it reflects a somewhat smaller base of lexical personality research. Thus the overall fit is anticipated to be similar to the Big Six. The Big Two is a parsimonious model drawing on a more diverse range of cultures than the Big Six (Saucier, Thalmayer et al, 2014). Thus, it might be anticipated to demonstrate better cross-cultural measurement invariance, particularly in non-western settings. The comparative fit of the models, however, is not the purpose of the current study. Because the three models are all derived from a Big Six measure, the Big Five and Big Two begin such a comparison at a disadvantage. Our main purpose in including additional models derived from a Big Six measure is to explore the relative fit of items and thereby refine measures of the three models for the use of researchers who collect survey data for cross cultural comparisons.

#### Method

#### **Participants**

Survey of World Views data included 8,883 participants from 33 countries. In the current study, several exclusion criteria were applied prior to analysis. Participants were eliminated if they were not students; if more than 10% of a participant's 36QB6 responses were missing; if that participant's standard deviation for 36QB6 items was below .50 (to cull those who tended to give the

same response for every item); or if they were one of a few cases judged to be a likely duplicate responder. Participants were also excluded if they were very extreme and consistent with respect to yea-saying or nay-saying in the full questionnaire. Finally, countries were excluded if the remaining sample was smaller than 150 participants. Criteria for even stricter exclusions were applied only as a last resort in individual country-samples where problems with analysis convergence were encountered (noted in text).

Table 1 displays sample sizes per country for the 7,378 participants from 26 nations included in the current analyses. Average age in the samples ranged from 19.8 in the Philippines and Ukraine to almost 24 in Kenya, Ethiopia, and Argentina, and 24.5 in Tanzania (average sample mean age = 21.7, SD = 1.28). In four of the 26 nations men were the majority of participants (Bangladesh [22% female], Ethiopia [28%], Tanzania [31%], and Kenya [34%]). In the remaining 22 nations, women were the majority, with the highest percentages in Thailand (75%), Brazil (78%) and Poland (89%; average across samples = 59% female).

The country samples were grouped into three subsets to facilitate model respecification and testing. Selection of countries into these subsets was made prior to this study by the second author, such that each has a high N and represents all major parts of the world in a similar way. Splitting the countries into three sets enabled us to: a) test *a priori* structural models, b) empirically derive models that might achieve a better fit, and c) rigorously test those models to establish their generalizability and usefulness to investigators in future studies. We were able to use the first set of countries as a derivation sample, in which a model is derived and optimized, and the second set of countries as a cross-validation sample, in which the fit of the optimized model is interpreted as a realistic estimate of the generalizability of the optimized model to other samples (Wiggins, 1973). This procedure could then be repeated to further refine the model for cross-validation in the third set of countries. Cross-validation is recommended for regressionbased procedures when the sample is large enough (Horst, 1966; Wiggins, 1973). Each of our sets of countries had over 2,000 cases, which seems a sufficient sample size for empirically deriving or for testing a model.

#### Procedure

Country selection attempted to represent the world, in terms of demographic footprint and economic impact. The 33 sampled countries have aggregated populations amounting to 67.3 percent (4.7 billion) of the world's population; when the gross domestic product of these 33 countries is aggregated, the total makes up 76.2 percent of the gross aggregate domestic products of all countries in the world (Central Intelligence Agency, 2012).

Cooperating faculty from diverse fields publicized the study to students at their own higher-education institutions. Data were collected online in 2012 via a U.S. server platform, with compensation handled via Western Union or Amazon gift coupons. See Saucier, Kenner et al (2014) for details about data collection and the full sample. Use of college students enabled standardized online administration and minimized between-population differences in level of education.

#### Materials

The current study used 40 total QB6 items (the 36QB6 and, appended at the end, four QB6 items from longer versions; Saucier, 2009; see Table 2). Questionnaires for participants in Canada, England, India, Kenya, Singapore, and the United States were in English. Participants in other countries used items translated into Chinese (China, Taiwan), 'new world' Spanish (Peru, Argentina), Castilian Spanish (Spain), Arabic (Morocco), Kiswahili (Tanzania), Amharic (Ethiopia), Portuguese (Brazil), or German, Polish, Ukrainian, Greek, Turkish, Japanese, Thai, Malay, Nepali, Bengali, or Filipino/Tagalog. In all cases back-translation was used, with at least two translators working independently.

In addition to the QB6, a Big Five model and indicators of the Big Two, Social Self-

regulation and Dynamism, were analyzed. The initial Big Five included all items from the 36QB6, with Agreeableness (A) and Honesty/Propriety (H) items collapsed into a single scale. This conceptualization of an A/H scale emphasizes H more than some Big Five A scales (there is greater emphasis here on patience and on a lack of hostility or taking advantage of others, and less on being actively kind). However, DeRaad et al (2010) argue that the Honesty dimension should rightly be considered part of Agreeableness, in part based on their interpretation of Ashton and Lee (2007) that two of the six facets of NEO-PI-R A are Honesty-related. As evident in Thalmayer et al (2011, supplemental materials), the other scales of the 48QB6 (a slightly longer version of the 36QB6) correlate highly with analogous BFI (John, Donahue, & Kentle, 1991) and NEO-FFI (Costa & McCrae, 1992) Big Five scales. The convergent correlations were .80 to .81 for Conscientiousness, .76 to .82 for Resiliency and Emotional Stability versus Neuroticism, .68 to .70 for Extraversion, .63 to .74 for Openness and Originality. The highest divergent correlation was only .36 (BFI Agreeableness with QB6 Extraversion).

Ten-item Big Two scales (see Table 3) were developed from the 40 personality items available in the Study of World Views data using the following procedure:

- The 40 items were correlated in the Eugene Springfield Community Sample (N = 453) with Big Two adjective markers from Saucier, Thalmayer et al (2014).
  - a. For Social Self-Regulation (S) terms included: Honest, kind, generous, gentle, good, obedient, respectful, diligent, responsible and (reverse keyed) selfish.
  - b. For Dynamism (D), terms included: Active, brave, bold, lively and (reverse keyed) timid, weak, and shy.
- 2. A reduced set with at least (roughly) double the loading on the primary versus secondary factor (relatively univocal) were retained.
- 3. An EFA indicated items with low loadings that could be dropped, resulting in 13 S and 12

D items retained.

 An EFA with the 25 items in a college student sample (N=225) indicated four items with low loadings and one overly redundant item for removal, leading to scales with 10 items each.

#### Analyses

The 36QB6, 36-item Big Five, and the 20-item Big Two were initially tested individually in each of the set 1 countries using confirmatory factor analysis in Mplus version 7. The set including the United States was chosen as set 1 to facilitate comparison with fit in the country most influential in creating the QB6 model. Comparisons were made to standard benchmarks per Hu and Bentler (1999) and to domain specific benchmarks fit statistics (those reported by Hopwood and Donnellan [2010]), as detailed above.

Measurement invariance was then tested in four stages (as described above, and following Muthén and Muthén [2012]):

- 1. Configural invariance: Factor means fixed at zero in all groups but factor loadings and other parameters allowed to vary.
- Factorial/metric invariance: Adds constraint of equal factor loadings across samples to above.
- Scalar invariance: Adds constraint of equal intercepts across groups to above; factor means fixed at zero in one group and free in others.
- 4. Full/strict invariance: Adds constraint of equal error variances to all above constraints. According to Cheung and Rensvold (2002) and Kline (2011), change in CFI between one level and the next of more than .01 indicates that the null hypothesis of invariance should not be rejected (in other words, fit may be worse at the stricter level). Fit at all levels is reported,

however, regardless of whether the criterion is met, for relative comparisons.

Next, items were removed from the 36QB6, the 36-item Big Five, and the 20-item Big Two based on review of standardized loadings and intercepts across the set 1 countries (for this step, all 40 available items were considered for the QB6, and 39 of the 40 (excluding additional H item "I stick to the rules", which was believed to be more similar to Big Five Conscientiousness than Agreeableness content). The refined, provisional models were then tested individually in each set 2 nation, and again standardized loadings and intercept variation across countries were perused to identify items for removal. The final refined models were then tested individually and for measurement invariance in the set 3 nations. For comparison purposes, the original, full-length models were also run in the set 3 nations.

#### Results

#### **Questionnaire Big Six**

The fit of the 36QB6 in the first set of nations is reported in Table 4. The model converged in all, and fit was similar across countries. Notably, fit was not better in the United States than in other cultural groups. The fit of the QB6 across nations did not meet standard benchmarks for good fit in any nation, but it was similar to or better than that reported by Hopwood and Donnellan (2010) for broad personality inventories developed and tested within North American samples. Thus, fit can be said to have met domain specific benchmarks in most of the set 1 nations.

The results of testing the four levels of measurement invariance of the 36-QB6 in set 1 are also reported in Table 4. These indicate little change in fit between configural and factorial invariance, but some decline in fit between factorial and scalar levels. Between all levels, however, the change was larger than the .01 criteria proposed by Cheung and Rensvold (2002) to indicate non-significant change. The last stage, full measurement invariance, had to be tested excluding Kenya, due to difficulty with convergence for the Kenyan group (not positive definite) in this test only. For comparison purposes, the other three levels are additionally reported excluding Kenya.

Next, in hopes of revising for maximum cross-cultural fit, we consulted modification indices. However, these were difficult to act on, because of inconsistency in indications across groups and lack of interpretability. Instead, we considered the item pool systematically, in terms of indices of item difficulty and discrimination. The four additional QB6 items (two Extraversion [E], one Honesty/Propriety [H] and one Originality [O]) were added to the 36 original items, and the 40 items were perused for differential item functioning across set 1 countries. Items that fit relatively poorly cross-culturally were identified in terms of: (a) number of countries for which standardized loadings of the item on intended factor was lower than .25; and (b) high standard deviation in intercepts, indicating greater relative variation in endorsement (difficulty) across nations. On this basis, seven items (one each from Conscientiousness [C], H, and Agreeableness [A], and two each from O and E) were removed from the group (see Table 2)<sup>1</sup>.

The provisional 33-item QB6 was then tested using CFA in the second set of nations (see Table 5). Inspection of item standardized loadings and intercepts in this set, and an effort to maintain balanced keying, led to further removal of one item each from the E, O, H, and Resiliency scales. (Because the model did not converge in the data from Morocco or Tanzania, standardized loadings from these groups were not available.) This led to a refined 30-item version (henceforth 30QB6), with exactly five items on each scale.

The fit of the 30QB6 and the progressive measurement invariance analyses in set 3

<sup>&</sup>lt;sup>1</sup> One of the E item ("I talk a lot") was removed not due to poor fit, but to avoid redundancy with an added item ("I don't talk a lot"); the former was chosen for removal in the interest of balanced keying.

nations are reported in Table 6. Due to some difficulties with convergence in the data from Ethiopia, more stringent data selection criteria were used in that group,<sup>2</sup> and one item with especially poor fit was excluded (R36: "I rarely worry"). Fit indices indicate better fit than for the 36QB6 in set 1, and similar fit to domain specific benchmarks. For direct comparison, progressive measurement invariance analyses for the 36QB6 in this set of nations are also included. (At the scalar and full levels, the analysis could not be reported due to difficulties with the Ethiopia set, and analyses are reported excluding Ethiopia. Comparable analyses of the 30QB6 were added.) The comparison favors the 30QB6 over the 36QB6, as hypothesized, particularly (in some cases only) in terms of CFI and TLI.

#### The Big Five

The fit of the 36-item Big Five in the first set of nations is reported in Table 7. The model converged in all, and fit was similar across countries. Again, fit was not better in the United States than for other cultural groups. As for the 36QB6, fit across nations failed to meet standard benchmarks, while generally meeting domain specific benchmarks. The results of testing the four levels of measurement invariance again indicate little change in fit between configural and factorial invariance, but some decline in fit in scalar and full levels. Fit is slightly better for the Big Five than for the Big Six model in terms of CFI and TLI. Again, CFI difference between

<sup>&</sup>lt;sup>2</sup> This set of criteria was developed independently of and prior to the present study, including elimination of cases that might be problematic from a data-quality standpoint. Cases were excluded if any two of the following conditions were met: full (over 300-item) questionnaire completed in under 20 minutes, very low variance in responses across questionnaire, tendency to perseverate (give highly similar responses to adjacent items) across parts of the long questionnaire, high possibility based on cluster analysis of cases that the case was either random in responding or non-independent of another case, and having a response profile (across personality or other items) that was negatively correlated with the typical response profile. The criteria were applied in a conservative way, resulting in a set with N=305, and in a more liberal way, resulting in a set with N = 283. The largest set in which analyses would converge was used. The set used for a set of analyses can be seen in N size in top of Tables 6, 9, and 12.

levels of invariance was always greater than .01, indicating that more stringent models fit significantly more poorly.

To revise for maximum cross-cultural fit, three of the four additional QB6 items (two Extraversion [E] and one Originality [O]) were added to the 36 original items, and CFA results for this 39-item version in each set 1 country were perused as described above for the QB6. Here too, seven items (one each from Conscientiousness [C] and O, two from E, and three from Agreeableness/Honesty [A/H]) were removed from the group. This provisional 32-item Big Five measure was then tested using CFA in the second set of nations (see Table 8). Inspection of standardized loadings and intercepts, and an effort to maintain balanced keying and scales of similar length, next led to removal of one item each from the E, O, and Resiliency scales, and all four remaining H items from the A/H scale. (Because the model did not converge in the data from Morocco or Nepal, standardized loadings from these groups were not available.) This led to a 25-item Big Five, with five items on each scale.

The fit of the 25-item Big Five measure (henceforth 25QB6) and the progressive measurement invariance analyses in set 3 nations are reported in Table 9. Fit indices indicate better fit than the 36-item version in set 1.

#### The Big Two

The 20-item Big Two measure fit slightly less well in the set 1 data than did the 36QB6 (see Table 10). Again, change in CFI between levels of measurement invariance was greater than .01 in all cases. There was an especially substantial drop off in fit between factorial and scalar levels, indicating variation in scale means across groups (see Table 1 for scale means by nation).

As above, the standardized loadings of items on their factors and the standard deviation of item intercepts were consulted to identify cases of differential item functioning for removal. Balanced keying and maintenance of relatively even proportions of Big Six factor content in each of the Big Two were also considered. On the Social Self-Regulation scale, one H and two A items stood out as having particularly problematic loadings and intercepts. Three others (two H and one C item) had more minor indications of poor fit. Because the scale was made with three H and three C, but only two A items, however, the worst fitting H and C items were removed, and both A items were retained. Similarly, for Dynamism, because the scale was made from three E and O but only two Resiliency (R) items, E and O items were chosen for removal from the pool of those with most problems in fit (see Table 3).

This provisional 16-item Big Two measure was tested in the second set of nations (see Table 11). Again, standardized loadings and intercepts across countries were consulted to identify differential item functioning. For S, the poorest fit was observed for one item each from C, H, and A. Only the C item was removed in order to retain an adequate range of content. For D, the poorest fit was observed for one E and the two R items. In the interest of retaining the few remaining reverse keyed items, only one R item was removed.

The refined 14-item Big Two inventory (henceforth 14QB2) was tested in set 3 nations (see Table 12). Measurement invariance was still very poor at scalar and full levels. Compared to the 20-item Big Two in the same datasets, fit was slightly improved at all levels.<sup>3</sup>

<sup>&</sup>lt;sup>3</sup> To explore the role of response biases, an additional set of measurement invariance tests were conducted, with the average acquiescence tendency for each person removed. This average was calculated using 10 heterogeneous pairs of items with opposite meaning in the full data set (e.g. "I talk a lot and "I don't talk a lot"). The average response to the 20-items (which should logically be the midpoint of the response scale) was subtracted from each of that participant's item responses, for acquiescence-adjusted datasets centered around the participant's mean response to the pairs. For all models, results indicated better convergence (it was not necessary to exclude Ethiopia at any level for MI analyses in the QB6, for example, as it was in the non-adjusted data, and improved fit at the scalar and full levels. Overall fit, however was still quite poor at these levels. Results are available from the author. Although many factors make it difficult to determine the extent to which observed mean differences constitute true national differences versus response characteristics, this strategy for addressing the issue of response styles may be a promising future direction for cross-cultural survey research (which could be applicable in other kinds of group comparisons as well). Including matched pairs of forward- and reverse-keyed items allows for a quantatative assessment of a participants' tendency to yea- or nay-say.

Table 14 presents reliability values for the refined scales. The modest level of many of these indicators of internal consistency can be attributed in large part to the abbreviated nature of the scales. The scales are designed to capture the core of each factor in a maximally cross-culturally generalizable way, with the expectation that future psychometric work can effect an increase the number of items and can rebuild internal-consistency indices to more consistently adequate levels.

#### Discussion

In the current study, data from an unusually diverse group of nations was used to test the measurement invariance and cross cultural applicability of the 36QB6, a measure developed to improve on the cross-cultural validity of similar inventories. Big Five and the Big Two models were also tested. Additionally, the large dataset was used to refine more cross-culturally informed versions of all three measures. Refined (but still provisional) versions presented here are the 30QB6, the 25QB5, and the 14QB2.

Compared to standard fit indices (Hu & Bentler, 1999), the fit of the initial models in set 1 was marginal. While SRMR was under .09 for most, and RMSEA (which rewards parsimony) indicated close fit in some instances and adequate fit in many. in no case did CFI or TLI "incremental fit indices" meet the benchmark (.90 or above). This was anticipated due to analyzing an item-level, Likert-scale measure (Kline, 2011), with multiple factors, and because broad-bandwidth personality inventories of this nature consistently achieve poor fit, even when they demonstrate strong criterion validity (Hopwood & Donnellan, 2010). Fit in CFA may be fundamentally constrained for such inventories, given cross loadings and similarities in wording or life domain that logically result in correlated errors (Poortinga et al, 2002). For CFI (which assesses variance explained over the null model), low values were likely due to low standardized loadings of items on factors (low 'factor saturation'). Compared to fit indices reported by Hopwood and Donnellan (2010) for an array of personality inventories developed and tested and in North America, the results of CFA in the individual countries suggested similar levels of fit. This might be taken to indicate some level of cross-cultural fit in this particular field of study.

As predicted, neither the Big Six nor Big Five models had a clear advantage over the other. Comparing the refined 30- and 25-item versions, the Big Five had slightly better CFI and TLI values, but the QB6 had slightly better RMSEA values. SRMR values were mixed for CFA in individual countries, but slightly better for across- country analyses in the QB6.

In the tests of measurement invariance, more stringent levels in no cases met the criteria of less than .01 decrease in CFI. Thus, strictly speaking, the inventories do not meet criteria for measurement invariance. Looking at relative fit across levels, however, it can be seen that results suggest acceptable fit at the configural and factorial levels, provided that comparison is made to domain specific benchmarks. Fit at the configural level suggests that the same number of factors may work acceptably (again, by domain specific standards) across nations for the refined versions of all three inventories. At the factorial level, there is likewise indication that the items load in a rather consistent pattern on the same factors across nations, so that the factors would have similar interpretation.

The moderate (for the QB6 and Big Five) or dramatic (for the Big Two) drop-off in fit at the scalar level is not surprising given the divergences in scale means, observable in Table 1. It appears that while the models all work more or less adequately (by domain specific benchmarks) at basic levels, to explain the number of latent variables present for the items, and which items relate to which scales, they do not currently offer any basis for mean comparisons across cultures. Such comparisons, of course, are problematic for many reasons. Cultural diversity leads to challenges in scale translation because in many cases the same concepts (pride, insights, taking risks) convey different emotional or evaluative tones. Highly differential levels of familiarity with survey tasks can lead to differences in use of the scale options. Reference group effects can also affect responses and scores (Heine, 2012; Heine, Lehman, Peng, & Greenholz, 2002) – likely the case with Conscientiousness in the current data, where the lowest mean score was observed in Japan, a place hardly known for a lackadaisical, impulsive way of life (scores in Tanzania were two standard deviations higher). Consistent cultural differences in the amount of variance observed in trait scales, with Europeans expressing the greatest within-group variation and East Asians and Africans the least, have also been reported (McCrae, 2002). This may be because in individualistic, as opposed to collectivistic cultures, more diversity may be expressed and given importance (McCrae, 2002). It may also be due to response styles driven by similar cultural forces. For example, in East Asia there is more tendency toward middle responding (McCrae, 2002).

The difficulty of fitting the QB6 model in data from Africa (Morocco, Ethiopia, and Tanzania) provides an excellent illustration of the effect of population selection in developing a model. The Big Five was initially developed in a small range of nations, principally the United States (e.g. Goldberg, 1990), with crucial early confirmation in the Netherlands and Germany. The Big Six drew on data from a larger and more diverse group of countries (Ashton et al, 2004), but this group did not include any from the African continent. None of the models were developed using data from South America, either, but this has likely been less consequential, since South American countries use European languages and their populations are partly a European diaspora. The Big Two model conceptualized here, on the other hand, was developed based on data from nine nations, two of them African (Saucier, Thalmayer et al, 2014). And in the present study, the Big Two model fit just as well in data from Africa as it did in nations from other regions. It is of note that the final QB2 version fit best first in Ethiopia, secondly in Bangladesh, and worst in England. The Big Two is, if anything, strongest in the 'global south', although the differences are small, and the model works nearly as well in the 'global north'. Because of a dearth of lexical studies there, we don't really know what indigenous five- and six-factor models would be in the 'global south'; they may have their own replicable patterns e.g. an alternative 'southern Big Five' or Six.

Overall, however, it cannot be said that the QB2 fit better than the QB6 or QB5 – it simply fit more evenly across contexts, applying in a more trouble-free manner in more places. As items for this measure were chosen only from QB6 items, this particular measure of the model started at a disadvantage – few of the core Big Two items identified in the last table in Saucier, Thalmayer et al (2014) were available in the pool. Such core items would tend to be interstitial to Honesty, Agreeableness, and Conscientiousness, rather than representing one of these factors exclusively as in the QB6 item pool. The Big Two measure should thus especially be seen as in an early stage of development. Future cross-cultural surveys would ideally draw on a wider selection of items, hewing closer to the content in the adjectives identified in Saucier, Thalmayer et al (2014); we are not however advocating using actual adjectives as measures of the Big Two, since adjectives can be especially difficult to faithfully translate.

The QB6 is likewise still under development. Even the refined version presented here is not intended as a final, superior measure of the Big Five or Six, but as an intermediate iteration based on a large, interesting pool of items (IPIP; Goldberg et al, 2006). The current study applies a cross-cultural generalizability criterion to a relatively early stage of inventory development, with an eye toward creating an inventory and a model that is more culturally de-

#### THE QUESTIONNAIRE BIG SIX IN 26 NATIONS

centered. The results are informative as to which items translate more readily and comparably, leading to more consistent intercepts and factor loadings. A limitation of the current study is the restriction of participation to college students. While this facilitated cross-cultural comparison by holding age, literacy, and education-level relatively constant, it does limit our ability to generalize to the entire populations from which our samples were drawn.

A limitation specific to the Big Five and Big Two measures is that the QB6 items used were not chosen with the measurement of these models in mind. This was particularly problematic for the Big Two – the refined measures presented here are shorter and include fewer core-content items than would be ideal to cover the two broad dimensions. For the 25QB5, this limitation is specific to the Agreeableness domain, which here lacks some of the kindness and warmth content often emphasized in Big Five measures. It is our hope, however, that developing and validating Big Five and Big Two measures from this set of items will allow researchers, regardless of preferred model, to make use of the translations of the personality items used in the Survey of World Views. Translations of items on these inventories, now available in 31 languages, represent a significant cooperative effort on the part of translators, psychologists, and linguists around the globe. We hope to facilitate cross-cultural research by making these items and scales freely available to other researchers. While the scientists developing models of personality have historically come from a restricted range of nations, there is increasing awareness that broadening our scope of interest can improve the replicability, generalizability, and quality of our results. The measures presented here should ideally contribute to the long-term goal of understanding human personality in ways that transcend a single place and time.

#### References

- Ashton, M. C., Lee, K., Perugini, M., Szarota, P. De Vries, R. E., Di Blas, L., Boies, K., & De Raad, B. (2004). A six-factor structure of personality-descriptive adjectives: Solutions from psycholexical studies in seven languages. *Journal of Personality and Social Psychology*, 86, 356-366.
- Ashton, M. C., & Lee, K. (2007). Empirical, theoretical, and practical advantages of the HEXACO model of personality structure. *Personality and Social Psychology Review*, 11, 150–166. doi:10.1177/1088868306294907
- Central Intelligence Agency (2012). The World Factbook. (Retrieved 12/11/2012 from https:// www.cia.gov/library/publications/the-world- factbook/rankorder/2119rank.html )
- Cheung, G.W., & Rensvold R. B. (2002). Evaluating goodness-of-fit indexes for testing measurement invariance. *Structural Equation Modeling*, *9*, 233–255.
- Church, A. T., Alvarez, J. M., Mai, N. T., French, B. F., Katigbak, M. S., & Ortiz, F. A. (2011). Are cross-cultural comparisons of personality profiles meaningful? Differential item and facet functioning in the Revised NEO Personality Inventory. *Journal of Personality and Social Psychology, 101*, 1068-89.
- Costa, P. T., Jr., & McCrae, R. R. (1992). Revised NEO Personality Inventory (NEO-PI-R) and NEO Five-Factor Inventory (NEO-FFI) professional manual. Odessa, FL: Psychological Assessment Resources.
- Cramer, A. O. J., van, S. S., Noordhof, A., Borsboom, D., Wichers, M., Geschwind, N., Aggen, S. H., ... Borsboom, D. (2012). Dimensions of normal personality as networks in search of equilibrium: You can't like parties if you don't like people. *European Journal of Personality, 26*, 414-431.

- De Raad, B., Barelds, D. P., Levert, E., Ostendorf, F., Mlacić, B., Di, B. L., Hrebicková, M., ... Katigbak, M. S. (2010). Only three factors of personality description are fully replicable across languages: a comparison of 14 trait taxonomies. *Journal of Personality and Social Psychology*, 98, 160-73.
- De Raad, B., DiBlas, L., & Perugini, M. (1998). Two independently constructed Italian trait taxonomies: Comparisons among Italian and between Italian and Germanic languages. *European Journal of Personality*, *12*, 19–41.
- De Raad, B., Henriks, A.A. J., & Hofstee, W. K. B. (1992) Towards a refined structure of personality traits. *European Journal of Personality*, *6*, 301-319.
- DeYoung, C. G. (2006). Higher-order factors of the Big Five in a multi-informant sample. Journal of Personality and Social Psychology, 91, 1138-1151.
- Digman, J. M. (1997). Higher order factors of the Big Five. *Journal of Personality and Social Psychology*, 73, 1246-1256.
- Goldberg, L. R. (1981). Language and individual differences: The search for universals in personality lexicons. In L. Wheeler (Ed.), *Review of Personality and Social Psychology: Vol. 2* (pp. 141-165). Beverly Hills, CA: Sage.
- Goldberg, L. R. (1990). An alternative "Description of personality": The Big-Five factor structure. *Journal of Personality and Social Psychology*, *59*, 1216-1229.
- Goldberg, L. R., Johnson, J. A., Eber, H. W., Hogan, R., Ashton, M. C., Cloninger, C. R., &
  Gough, H. C. (2006). The International Personality Item Pool and the future of public
  domain personality measures. *Journal of Research in Personality*, 40, 84-96.
- Heine, S. J. (2012) Cultural psychology. NY: Norton.
- Heine, S. J., Lehman, D. R., Peng, K., & Greenholtz, J. (2002). What's wrong with cross-cultural comparisons of subjective Likert scales? The reference-group effect. *Journal of*

Personality and Social Psychology, 82, 903-918.

- Hopwood, C. J., & Donnellan, M. B. (2010). How should the internal structure of personality inventories be evaluated? *Personality and Social Psychology Review*, 14, 332-346. doi: 10.1177/1088868310361240
- Horst, P. (1966). An overview of the essentials of multivariate analysis methods. In R.B. Cattell (Ed.), *Handbook of Multivariate Experimental Psychology* (pp. 129-152). Chicago, IL: Rand McNally & Company.
- Hu, L., & Bentler, P. M. (1999). Cutoff criteria for fit indexes in covariance structure analysis:Conventional criteria versus new alternatives. *Structural Equation Modeling*, *6*, 1-55.
- John, O. P., Donahue, E. M., & Kentle, R. L. (1991). *The Big Five Inventory* (Versions 4a and 54). Berkeley, CA: University of California, Berkeley, Institute of Personality and Social Research.
- John, O. P. & Srivastava, S. (1999). The Big Five trait taxonomy: History, measurement, and theoretical perspectives. In Pervin, L. A. & John, O. P. (Eds.), *Handbook of Personality: Theory and Research*. NY: The Guilford Press.
- Kline, R. B. (2011). *Principles and practice of structural equation modeling*. New York: Guilford Press.
- Lee, K., & Ashton, M. C. (2004). Psychometric properties of the HEXACO Personality Inventory. *Multivariate Behavioral Research, 39, 329-358*.
- Marsh, H. W., Hau, K. T., & Wen, Z. (2004). In search of golden rules: Comment on hypothesis-testing approaches to setting cutoff values for fit indices and dangers in overgeneralizing Hu and Bentler's (1999) findings. *Structural Equation Modeling*, *11*, 320-341.

- McCrae, R.R. (2002) NEO-PI-R data from 36 cultures: further intercultural comparisons. In R.
  R. McCrae, & J. Allik (Eds.), *The Five Factor model of personality across cultures* (pp. 105-125). New York: Kluwer Academic/Plenum Publishers.
- McCrae, R. R. & Costa, P. T. (1997) Personality trait structure as a human universal. *American Psychologist*, 52, 509-516.
- Muthén, L. K. & Muthén, B. O. (1998-2012) *Mplus user's guide* (7th ed.). Los Angeles, CA: Muthén & Muthén.
- Nye, C. D., Roberts, B. W., Saucier, G., & Zhou, X. (2008). Testing the measurement equivalence of personality adjective items across cultures. *Journal of Research in Personality*, 42, 1524-1536.
- Ostendorf, F. (1990). Sprache und Personlichkeitstrucktur: Zur Validität des Fünf-Faktoren-Modells der Persönlichkeit. [Language and personality structure: Towards the validity of the Five-Factor model of personality] Regensburg, Germany: Roderer.
- Ozer, D. J., & Benet-Martínez, V. (2006). Personality and the prediction of consequential outcomes. *Annual Review of Psychology*, *57*, 401-21.
- Poortinga, Van de Vijver, F. & Van Helmert, B. (2002). Cross-cultural equivalence of the Big
  Five: A tentative interpretation of the evidence. In McCrae, R. R. & Allik, J., (Ed.s) *The Five Factor model of personality across cultures*. New York: Kluwer Academic/Plenum
  Publishers.
- Saucier, G. (1994). Mini-markers: A brief version of Goldberg's unipolar Big Five markers. Journal of Personality Assessment, 63, 506–516. doi:10.1207/s15327752jpa6303\_8
- Saucier, G. (2009). Recurrent personality dimensions in inclusive lexical studies: Indications for a Big Six structure. *Journal of Personality*, 77, 1577-1614.

- Saucier, G., Georgiades, S., Tsaousis, I., & Goldberg, L. R. (2005). The factor structure of Greek personality adjectives. *Journal of Personality and Social Psychology*, *88*, 856–875.
- Saucier, G., & Goldberg, L. R. (1996). The language of personality: Lexical perspectives on the five-factor model. In J. S. Wiggins (Ed.), *The five-factor model of personality: Theoretical perspectives* (pp. 21-50). New York: Guilford.
- Saucier, G., Kenner, J., Bou Malham, P., Iurino, K., Chen, Z., Thalmayer, A.G., ..., Kovaleva,
  A. (2014). *Cross-cultural differences in a global "Survey of World Views"*. Submitted for publication.
- Saucier, G., Thalmayer, A.G., & Bel-Bahar, T. (in press) Personality descriptors ubiquitous across 12 languages. *Journal of Personality and Social Psychology*.
- Saucier, G., Thalmayer, A.G., Payne, D. L., Carlson, R., Sanogo, L., Ole-Kotikash, L., Church, A.T., Katigbak, M. S., Szarota, P., Szirmák, Z., & Zhou, X. (2014). A basic bivariate structure of personality attributes evident across nine languages. *Journal of Personality*, 82,1-14.
- Szirmak, Z., & De Raad, B. (1994). Taxonomy and structure of Hungarian personality traits. *European Journal of Personality*, *8*, 95–117.
- Thalmayer, A. G. Saucier, G. & Eigenhuis, A. (2011). The comparative validity of brief- to medium–length Big Five and big six personality inventories. *Psychological Assessment,* 23, 995 1009. doi: 10.1037/a0024165
- Vandenberg, R. J., & Lance, C. E. (2000). A review and synthesis of the measurement invariance literature: Suggestions, practices, and recommendations for organizational research. *Organization Research Methods*, 3, 4-70.

Wiggins, J.S. (1973). Personality and prediction: Principles of personality assessment.

Philippines: Addison-Wesley Publishing Company, Inc.

Zhou, X., Saucier, G., Gao, D., & Liu, J. (2009). The factor structure of Chinese personality descriptors. *Journal of Personality*, 77, 363–400.

# Sample Sizes and Means and Standard Deviations of Scales for the 26 Countries, Grouped by Region

						36	6QB6 M	lean (S	<u>D)</u>					20QB2 Mean (SD)			
Country/Region	Ν		С		H		А	]	R	]	E	(	0		S	Ι	)
Africa (sub-Saha	ran)																
Tanzania	209	4.14	(.73)	3.88	(.66)	3.01	(.59)	3.26	(.71)	3.35	(.56)	3.27	(.45)	3.95	(.60)	3.41	(.41)
Kenya	237	4.09	(.62)	3.87	(.63)	2.97	(.64)	3.20	(.68)	3.63	(.57)	3.23	(.55)	3.80	(.53)	3.31	(.47)
Ethiopia	331	3.79	(.69)	3.79	(.65)	3.06	(.62)	3.19	(.68)	3.09	(.57)	3.00	(.53)	3.78	(.61)	3.11	(.37)
North Africa/Mid	ldle Ea	st															
Morocco	342	3.31	(.78)	3.41	(.85)	2.93	(.60)	3.02	(.66)	3.20	(.66)	3.05	(.49)	3.34	(.67)	3.09	(.46)
Turkey	396	3.62	(.73)	3.65	(.66)	2.76	(.71)	2.91	(.82)	3.68	(.62)	3.48	(.56)	3.53	(.54)	3.37	(.51)
South Asia																	
Bangladesh	242	3.67	(.67)	3.55	(.56)	2.87	(.62)	2.92	(.76)	3.27	(.63)	3.02	(.58)	3.54	(.51)	3.12	(.47)
India	333	3.47	(.63)	3.58	(.62)	2.77	(.71)	2.96	(.71)	3.49	(.65)	3.20	(.58)	3.44	(.52)	3.22	(.53)
Nepal	314	3.73	(.62)	3.79	(.58)	2.73	(.56)	2.77	(.75)	3.58	(.65)	3.00	(.44)	3.63	(.47)	3.13	(.41)
Southeast Asia																	
Malaysia	299	3.98	(.62)	3.60	(.55)	2.94	(.55)	2.81	(.66)	3.60	(.58)	3.21	(.53)	3.79	(.45)	3.21	(.44)
Philippines	362	3.67	(.71)	3.85	(.62)	3.05	(.61)	2.82	(.64)	3.78	(.64)	3.54	(.60)	3.80	(.52)	3.42	(.48)
Thailand	313	3.59	(.61)	3.58	(.59)	2.86	(.63)	2.69	(.74)	3.63	(.65)	3.15	(.51)	3.52	(.45)	3.22	(.44)
Singapore	280	3.46	(.60)	3.46	(.61)	2.76	(.64)	2.90	(.74)	3.56	(.59)	3.38	(.63)	3.39	(.49)	3.21	(.49)
East Asia																	
Mainland China	285	3.60	(.58)	3.66	(.56)	2.86	(.60)	2.91	(.67)	3.59	(.55)	3.33	(.57)	3.58	(.44)	3.17	(.46)
Taiwan	352	3.46	(.61)	3.34	(.57)	2.96	(.66)	2.82	(.73)	3.73	(.65)	3.26	(.58)	3.37	(.45)	3.23	(.52)
Japan	366	2.74	(.70)	3.63	(.69)	3.03	(.70)	2.44	(.76)	3.67	(.72)	2.97	(.71)	3.19	(.50)	2.87	(.61)
East/Southeast E																	
Ukraine	210	3.59	(.68)	3.72	(.65)	2.80	(.68)	3.12	(.82)	3.77	(.66)	3.57	(.61)	3.55	(.54)	3.36	(.48)
Poland	223	3.34	(.77)	3.57	(.65)	2.83	(.80)	2.60	(.84)	3.62	(.75)	3.94	(.54)	3.38	(.58)	3.44	(.57)
Greece	228	3.50	(.73)	3.85	(.69)	2.89	(.69)	2.78	(.76)	3.89	(.59)	3.18	(.58)	3.57	(.55)	3.26	(.44)
Western Europe																	
Spain	322	3.67	(.69)	3.61	(.68)	3.00	(.67)	3.04	(.66)	3.80	(.67)	3.56	(.60)	3.63	(.55)	3.44	(.45)
Germany	306	3.49	(.71)	3.73	(.69)	2.89	(.69)	3.04	(.74)	3.80	(.65)	3.56	(.54)	3.54	(.51)	3.49	(.49)
United Kingdom	164	3.40	(.69)	3.55	(.72)	3.00	(.74)	2.83	(.88)	3.85	(.63)	3.50	(.55)	3.46	(.57)	3.41	(.43)

### North America

Canada	200	3.52	(.63)	3.58	(.75)	3.08	(.71)	2.87	(.84	3.82	(.72)	3.65	(.59	3.55	(.53)	3.47	(.51)
United States	391	3.58	(.64)	3.57	(.68)	3.06	(.63)	2.99	(.79)	3.72	(.63)	3.57	(.60)	3.54	(.52)	3.43	(.45)
Latin America																	
Peru	266	3.37	(.75)	3.62	(.58)	2.83	(.67)	2.97	(.67)	3.80	(.63)	3.47	(.55)	3.38	(.53)	3.45	(.49)
Argentina	214	3.69	(.69)	3.84	(.57)	2.69	(.68)	2.94	(.66)	3.85	(.66)	3.38	(.59)	3.56	(.50)	3.32	(.50)
Brazil	193	3.25	(.73)	3.98	(.67)	2.84	(.71)	2.69	(.67)	3.77	(.69)	3.68	(.59)	3.55	(.52)	3.22	(.54)
Total	7378	3.56	(.73)	3.65	(.66)	2.91	(.66)	2.90	(.75)	3.63	(.67)	3.34	(.62)	3.54	(.55)	3.28	(.51)
	· ·	тт	тт	/ / <b>D</b>	• ,		1 1	т	ת ו	.1.	T T		• •		• • • • •	/TC 1 (	C

*Note*. C = Conscientiousness, H = Honesty/Propriety, A = Agreeableness, R = Resiliency, E = Extraversion, O = Originality/Talent, S = Social Self Regulation, D = Dynamism.

NOTE: This is a corrected version 2016, There were two typos in the published version of this table, which made it unclear which HP and O items below in the final 30QB6 and 25QB5.

Initial QB6 Personality Items and Final Big Five and Big Six Inventories

Conscientiousness	Extraversion
1. I complete my duties as soon as possible.	3. I usually enjoy being with people.
7. I leave a mess in my room.	9. I reveal little about myself.
13. I like to plan ahead.	15. I laugh a lot.
19. I shirk my duties.	21. I don't think it's important to socialize
25. I like order. <sup>56</sup>	with others. <sup>5</sup>
<i>31. I waste my time.</i>	27. I talk a lot. <sup>56</sup>
	<i>33. I seldom joke around.<sup>56</sup></i>
	37. I am skilled in handling social situations. <sup>6</sup>
	40. I don't talk a lot.
Agreeableness	Originality
2. I hate waiting for anything.	4. I have difficulty understanding abstract
8. I am usually a patient person.	ideas.
14. I get angry easily.	10. I have a rich vocabulary. <sup>5</sup>
20. I am quick to correct others. <sup>56</sup>	16. I am considered to be a wise person.
26. I become frustrated and angry with people	22. I seldom experience sudden intuitive
when they don't live up to my expectations.	insights. <sup>56</sup>
32. I rarely show my anger.	28. I don't pride myself on being original. <sup>6</sup>
	34. I am an extraordinary person.
	39. I can handle a lot of information.
Honesty/Propriety	Resiliency
5. I take risks that could cause trouble for	6. I get stressed out easily.
<i>me</i> . <sup>56</sup>	12. I recover quickly from stress and illness.
11. I would never take things that aren't	18. I panic easily. <sup>6</sup>
mine. <sup>5</sup>	24. I am often worried by things I said or did.
17. I cannot imagine (that I would engage in)	30. I am afraid of many things.
lying or cheating. <sup>5</sup>	36. I rarely worry. <sup>5</sup>
23. I steal things. <sup>5</sup>	
29. I am not good at deceiving people. <sup>5</sup>	
<i>35. I like to do frightening things.</i> <sup>56</sup>	
38 L stick to the rules $5$	

<u>38. I stick to the rules.</u><sup>5</sup>

*Note*. Reverse keyed items italicized. Items 1 through 36 comprise the 36QB6, and the 36-item Big Five model tested. Items are available translated into 31 languages at:

http://psychometriglossia.uoregon.edu/

<sup>5</sup> Removed at first stage of revision process for Big Five.

<sup>5</sup> Removed at second stage of revision for Big Five. Unmarked items are included in the 25QB5.

<sup>6</sup> Removed at first stage of revision process for QB6.

<sup>6</sup>Removed at second stage of revision, QB6. Unmarked items are included in the 30QB6.

Big Two Items, with Big Six Scale Source Noted

Social Self-Regulation	Dynamism
C1. I complete my duties as soon as possible	O10. I have a rich vocabulary
A8. I am usually a patient person	E15. I laugh a lot
A14. I get angry easily	R30. I am afraid of many things
H17. I cannot imagine lying or cheating	O34. I am an extraordinary person
C19. I shirk my duties	E37. I am skilled at handling social situations
H23. I steal things	O39. I can handle a lot of information
H38. I stick to the rules	E40. I don't talk a lot
H5. I take risks that could cause trouble for	E9. I reveal little about myself*
me.*	O28. I don't pride myself on being original*
C7. I leave a mess in my room**	R36. I rarely worry**
C25. I like order*	

*Note*. Items denoted by number and Big Six domain, and italicized if reverse keyed. C = Conscientiousness, H = Honesty/Propriety, A = Agreeableness, R = Resiliency, E = Extractoristic (Talant

Extraversion, O = Originality/Talent.

\* Removed after examining model results in set 1 nations.

\*\* Removed after examining results in set 2 nations. Unmarked items were included in final, 14item version.

*Fit Indices of the 36-QB6 in Set 1 Nations, Individually, and for Progressively More Stringent Measurement Invariance Tests* 

Nation	Ν	χ2	df	CFI	TLI	AIC	RMSEA	SRMR
Argentina	214	1116.16	579	.582	.546	22401	.066	.082
Germany	306	1438.22	579	.646	.615	31294	.070	.087
Greece	228	1042.85	579	.678	.650	24324	.059*	.078
India	333	1285.45	579	.524	.482	37864	.061	.077
Kenya	237	1221.40	579	.513	.471	25434	.068	.087
Malaysia	299	1302.66	579	.585	.548	30407	.065	.087
Taiwan	352	1557.93	579	.630	.590	35370	.069	.083
Turkey	396	1594.55	579	.660	.630	41518	.067	.078
USA	391	1557.25	579	.622	.589	40963	.066	.078
			Measurer	nent invar	iance			
Configural	2756	12132.35	5223	.614	.581	289569	.066	.082
Factorial	2756	13082.41	5499	.576	.563	289967	.067	.094
Scalar	2756	17906.47	5739	.320	.328	294311	.083	.113
Full	2519	17998.42	5346	.237	.280	269818	.087	.145

*Note.* All adjusted  $\chi_2$  values p < .01. CFI = comparative fit index; TLI = Tucker-Lewis index; RMSEA = root mean square error of approximation; SRMSR = standardized root mean square residual.

1 Excluding Kenya.

\*Probability RMSEA is  $\leq .05$  was > .001

Nation	Ν	$\chi^2$	df	CFI	TLI	AIC	RMSEA	SRMR
Canada	200	986.12	480	.702	.672	18866	.073	.090
China	285	965.72	480	.684	.652	25827	.060*	.076
Nepal	314	1010.38	480	.651	.616	30135	.059*	.078
Peru	266	1015.51	480	.643	.607	25705	.065	.080
Spain	322	1235.26	480	.658	.624	30210	.070	.085
Thailand	313	1271.50	480	.572	.529	30089	.073	.086
Ukraine	210	813.037	480	.734	.708	19832	.057*	.082
			Meas	urement I	<u>nvariance</u>			
Configural <sub>1</sub>	1910	7297.53	3360	.659	.625	180665	.066	.082
Factorial <sub>2</sub>	2035	8858.91	4071	.607	.592	194324	.068	.097
Scalar <sub>2</sub>	2035	15258.44	3696	.376	.380	183580	.084	.113
Full <sub>2</sub>	2035	13152.63	4491	.289	.331	197778	.087	.143

The Provisional 33-item QB6 in Set 2 Nations

*Note*. Tanzania and Morocco were not included individually because the model did not converge in it (however, item means were used to infer comparability of intercepts). All adjusted  $\chi^2$  values p < .01. CFI = comparative fit index; TLI = Tucker-Lewis index; RMSEA = root mean square error of approximation; SRMSR = standardized root mean square residual.

\*Probability RMSEA is  $\leq .05$  was > .001

<sup>1</sup> Excluding Tanzania and Morocco due to non positive definite outcomes.

2 Excluding only Tanzania (Morocco included).

Nation	Ν	χ2	df	CFI	TLI	AIC	RMSEA	SRMR			
Bangladesh	242	854.69	390	.604	.558	22078	.070	.089			
Brazil	193	829.60	390	.616	.571	16949	.076	.090			
England	164	880.24	390	.599	.553	14304	.088	.096			
Ethiopia**	283	710.53	362	.654	.612	24612	.058*	.074			
Japan	366	1209.92	390	.641	.600	32952	.076	.084			
Philippines	362	1062.45	390	.674	.636	30421	.069	.082			
Poland	223	691.07	390	.799	.776	19040	.059*	.076			
Singapore	280	831.34	390	.712	.678	23408	.064	.075			
	Measurement invariance										
Configural	2113	7157.79	3127	.663	.625	184648	.070	.083			
Configural	1830	6375.21	2736	.667	.629	159156	.071	.084			
Factorial	2113	7983.97	3330	.611	.593	185068	.073	.103			
Scalar	2113	10725.67	3498	.396	.399	187473	.088	.123			
Scaları	1830	9109.78	3054	.445	.447	161254	.087	.117			
Full	2113	12287.32	3708	.281	.325	188153	.094	.174			
	Meas	urement inv	ariance	of the 360	QB6 (for	compariso	<u>n)</u>				
Configural	1830	9379.69	4053	.615	.581	192627	.071	.088			
Factorial	2113	11656.05	4884	.556	.541	224285	.072	.104			
Scaları	1830	13408.21	4449	.353	.359	195863	.088	.125			
Full	2113	17473.30	5346	.204	.250	229178	.093	.161			

The 30 QB6 in Set 3 Nations

*Note.* All adjusted  $\chi_2$  values p < .01. CFI = comparative fit index; TLI = Tucker-Lewis index; RMSEA = root mean square error of approximation; SRMSR = standardized root mean square residual.

\*Probability RMSEA is  $\leq .05$  was > .001

\*\*The 30-item version was non-positive definite in Ethiopia. Thus, the version tested here is 29 items, excluding item 36.

1 Excluding Ethiopia, in cases where analyses were non positive definite, or for comparison.

Fit Indices of Initial 36-item Big Five Measure in Set 1 Nations, Individually, and for Progressively More Stringent Measurement Invariance Tests

Nation	N	χ2	df	CFI	TLI	AIC	RMSEA	SRMR
Argentina	214	851.51	395	.601	.560	18785	.073	.084
Germany	306	1079.71	395	.655	.620	25868	.075	.085
Greece	228	730.84	395	.706	.676	20208	.061*	.079
India	333	903.71	395	.571	.528	31439	.062	.076
Kenya	237	877.54	395	.520	.471	21195	.072	.087
Malaysia	299	929.09	395	.631	.593	25019	.067	.087
Taiwan	352	1078.67	395	.682	.650	29301	.070	.079
Turkey	396	1090.41	395	.716	.687	34315	.067	.076
USA	391	1121.41	395	.649	.613	33820	.069	.078
			Measurer	nent inva	riance			
Configural	2756	8662.88	3555	.650	.614	239949	.068	.081
Factorial	2756	9984.41	3845	.579	.571	240691	.072	.095
Scalar	2756	13740.94	3995	.331	.345	244147	.089	.116
Full	2519	15680.01	4235	.215	.274	245606	.094	.145

*Note.* All adjusted  $\chi_2$  values p < .01. CFI = comparative fit index; TLI = Tucker-Lewis index; RMSEA = root mean square error of approximation; SRMSR = standardized root mean square residual.

\*Probability RMSEA is  $\leq .05$  was > .001

Nation	Ν	$\chi^2$	df	CFI	TLI	AIC	RMSEA	SRMR
Canada	200	1066.75	454	.617	.581	18481	.082	.101
China	285	1059.53	454	.600	.563	25292	.068	.083
Peru	266	973.19	454	.628	.594	24949	.066	.082
Spain	322	1300.23	454	.599	.562	29512	.076	.095
Tanzania	209	1070.50	454	.548	.506	19875	.081	.095
Thailand	313	1312.31	454	.516	.472	29399	.078	.090
Ukraine	210	893.58	454	.649	.617	19392	.068	.089
			Meas	urement i	nvariance			
Configural <sub>1</sub>	1805	7932.33	3188	.569	.531	167137	.076	.096
Factorial	2244	10901.16	4358	.492	.480	210175	.078	.110
Scalar <sub>2</sub>	1910	11246.03	3532	.299	.311	179286	.089	.120
Full	2244	16421.86	4814	.099	.165	214784	.098	.166

### The Provisional 32-item Big Five in Set 2 Nations

*Note.* Morocco and Nepal were not included individually because the model did not converge in either. However, items means were included with the analyses of intercepts, and the groups are included in MI analyses, below, except where noted. All adjusted  $\chi^2$  values p < .01. CFI = comparative fit index; TLI = Tucker-Lewis index; RMSEA = root mean square error of approximation; SRMSR = standardized root mean square residual.

<sup>1</sup> Excluding Morocco and Nepal, due to non positive definite results.

<sup>2</sup> Excluding Morocco and Tanzania, due to non positive definite results.

Nation	Ν	$\chi^2$	df	CFI	TLI	AIC	RMSEA	SRMR			
Bangladesh	242	618.78	265	.624	.575	18497	.074	.092			
Brazil	193	560.83	265	.665	.621	14381	.076	.086			
England	164	594.78	265	.655	.609	11925	.087	.095			
Ethiopia	305	553.28	265	.708	.669	22550	.060	.073			
Japan	366	820.67	265	.698	.658	27848	.076	.085			
Philippines	362	790.45	265	.685	.643	25699	.074	.083			
Poland	223	504.92	265	.817	.793	16187	.064	.078			
Singapore	280	530.47	265	.798	.771	19368	.060*	.067			
Measurement invariance											
Configural	2135	5436.09	2165	.670	.634	156827	.075	.090			
Configural <sup>1</sup>	1830	4882.81	1900	.665	.630	134277	.077	.093			
Factorial	2135	5716.92	2295	.654	.639	156848	.075	.103			
Scalar	2135	8584.95	2435	.379	.388	159436	.097	.137			
Full	2135	9736.67	2610	.280	.338	160238	.101	.162			
	Measure	ement invari	ance of 2	36-item v	ersion (fe	or comparis	<u>son)</u>				
Configural <sup>1</sup>	1830	11048.24	4155	.501	.470	213754	.078	.102			
Factorial	2135	12772.02	4924	.489	.477	227526	.077	.109			
Scalar	2135	20387.84	5040	.233	.248	231234	.093	.134			
Full	2135	18637.83	5393	.137	.194	232454	.096	.168			

### The 25QB5 in Set 3 Nations

*Note.* All adjusted  $\chi_2$  values p < .01. CFI = comparative fit index; TLI = Tucker-Lewis index; RMSEA = root mean square error of approximation; SRMSR = standardized root mean square residual.

\*Probability RMSEA is  $\leq .05$  was > .001

1 Excluding Ethiopia, in cases where analyses were non positive definite, or for comparison.

Nation	Ν	χ2	df	CFI	TLI	AIC	RMSEA	SRMR
Argentina	214	393.04	169	.567	.513	12352	.079	.081
Germany	306	710.089	169	.454	.386	17633	.102	.096
Greece	228	355.33	169	.602	.552	13466	.070	.077
India	333	442.86	169	.539	.482	21167	.070	.075
Kenya	237	445.16	169	.500	.437	14188	.083	.086
Malaysia	299	477.31	169	.572	.519	16699	.078	.080
Taiwan	352	691.53	169	.485	.421	19879	.094	.090
Turkey	396	797.78	169	.474	.409	23246	.097	.093
-		1	Measure	ment inva	riance			
Configural	2756	4873.95	1521	.512	.451	161443	.085	.085
Factorial	2756	5432.76	1681	.453	.444	161681	.085	.101
Scalar	2756	7959.065	1825	.106	.163	163920	.105	.128
Full	2756	8939.367	1985	.000	.108	159970	.108	.170

Table 10The 20-item QB2 in Set 1 Nations

*Note.* All adjusted  $\chi_2$  values p < .01. CFI = comparative fit index; TLI = Tucker-Lewis index; RMSEA = root mean square error of approximation; SRMSR = standardized root mean square residual.

Nation	Ν	$\chi^2$	df	CFI	TLI	AIC	RMSEA	SRMR
Canada	200	340.93	103	.482	.396	9379	.107	.100
China	285	312.40	103	.528	.450	12759	.084	.079
Morocco	342	314.40	103	.780	.743	17693	.077	.070
Nepal	314	304.21	103	.624	.562	14556	.079	.077
Peru	266	249.77	103	.637	.578	12375	.073*	.072
Spain	322	348.22	103	.642	.583	14460	.086	.074
Tanzania	209	247.94	103	.701	.652	9595	.082	.076
Thailand	313	380.87	103	.504	.423	14570	.093	.083
Ukraine	210	208.96	103	.681	.628	9501	.070*	.076
		]	Measure	ment Inv	ariance			
Configural	2461	2707.69	927	.634	.573	114887	.084	.078
Factorial	2461	3498.53	1039	.498	.486	115422	.092	.111
Scalar	2461	5650.94	1167	.078	.147	117351	.119	.150
Full	2461	6897.58	1295	.000	.039	118341	.126	.255

# The Provisional 16-Item QB2 in Set 2 Nations

*Note.* All adjusted  $\chi_2$  values p < .01. CFI = comparative fit index; TLI = Tucker-Lewis index; RMSEA = root mean square error of approximation; SRMSR = standardized root mean square residual.

\*Probability RMSEA is  $\leq .05$  was > .001

Nation	Ν	$\chi^2$	df	CFI	TLI	AIC	RMSEA	SRMR			
Bangladesh	242	211.84	76	.671	.606	10156	.086	.077			
Brazil	193	228.24	76	.525	.431	7756	.102	.092			
England	164	190.34	76	.597	.518	6647	.096	.088			
Ethiopia	331	152.87	76	.815	.779	13663	.055*	.056			
Japan	366	354.24	76	.587	.505	15453	.100	.084			
Philippines	362	291.86	76	.628	.555	13862	.089	.073			
Poland	223	244.32	76	.641	.570	8767	.100	.087			
Singapore	280	263.58	76	.583	.500	11012	.094	.083			
		<u>N</u>	<i>Aeasurem</i>	ent invar	iance						
Configural	2161	1937.29	608	.631	.559	87316	.090	.079			
Factorial	2161	2430.91	706	.552	.507	87613	.095	.112			
Scalar	2161	4307.55	790	.024	.101	89322	.128	.156			
Full	2161	5033.92	888	.000	.057	89582	.131	.284			
	Measurement invariance of 20QB2 (for comparison)										
Configural	2161	4182.45	1352	.542	.485	126757	.088	.088			
Factorial	2161	5109.62	1492	.415	.404	127404	.095	.120			
Scalar	2161	8226.68	1618	.000	004	130269	.123	.166			
Full	2161	9148.06	1758	.000	.034	130910	.125	.235			

### The 14QB2 in Set 3 Nations

*Note.* All adjusted  $\chi_2$  values p < .01. CFI = comparative fit index; TLI = Tucker-Lewis index; RMSEA = root mean square error of approximation; SRMSR = standardized root mean square residual.

\*Probability RMSEA is  $\leq .05$  was > .001

	30QB6						25QB5					14QB2	
	С	А	Е	0	R	HP	С	А	Е	0	R	S	D
C1	.57						.55					.40	
C7	.39						.40						
C13	.43						.42						
C19	.57						.56					.45	
C31	.57						.59						
A2		.31						.31					
A8		.50						.47				.33	
A14		.81						.85				.25	
A26		.40						.39					
A32		.58						.56					
E3			.54						.50				
E9			.39						.36				
E15			.50						.49				.32
E21			.47										
E37									.59				.69
E40			.60						.58				.28
O4				.36						.37			
O10				.66									.56
O16				.69						.61			
O28										.31			
O34				.45						.47			.40
O39				.59						.58			.56
R6					.74						.73		
R12					.42						.40		
R18											.69		
R24					.46						.42		
R30					.56						.59		.24
R36					.57								
H11						.47							
H17						.49						.39	
H23						.64						.50	
H29						.34							
H38						.51						.48	
А	.05						.05						
Е	.07	11					.12	05					
Ο	.16	05	.36				.23	.03	.58				
R	.25	.45	.27	.44			.25	.44	.41	.60			
HP	.39	.12	08	05	27								
D												.00	

Standardized Loadings and Inter-Factor Correlations (Factorial Invariance), Refined Inventories

	30QB6						25QB5*			14QB2	
	С	Н	А	R	Е	0	R	Е	0	S	D
Argentina	61	55	53	53	70	63	66	70	55	44	57
Bangladesh	53	62	57	60	36	53	68	36	49	60	38
Brazil	59	56	68	62	65	56	62	64	43	40	64
Canada	58	69	70	71	76	58	75	76	55	61	61
China	57	59	62	61	50	58	66	52	55	55	51
England	63	66	67	75	65	59	79	67	53	63	51
Ethiopia	57	54	50	39	37	27	65	27	08	60	15
Germany	64	65	70	70	70	63	72	70	51	54	62
Greece	64	63	60	65	62	63	70	56	48	50	54
India	50	50	60	49	58	53	56	60	34	46	59
Japan	63	59	65	69	68	63	67	69	59	51	65
Kenya	58	61	51	50	49	55	55	45	39	49	46
Malaysia	65	53	52	56	51	60	61	52	47	51	56
Morocco	46	76	25	39	27	39	47	23	-10	67	45
Nepal	55	60	46	62	57	52	66	49	21	54	45
Peru	71	50	58	54	66	59	62	64	56	47	57
Philippines	73	54	58	51	65	63	60	61	58	59	56
Poland	69	51	75	78	73	56	76	73	53	53	68
Singapore	60	62	66	70	63	63	73	67	55	53	63
Spain	64	60	64	58	64	68	62	63	59	57	62
Taiwan	65	56	69	72	68	58	72	72	56	55	62
Tanzania	69	53	43	42	49	32	65	39	08	60	47
Thailand	57	61	61	67	54	48	65	52	31	53	48
Turkey	70	63	70	70	58	62	72	63	58	56	56
Ukraine	59	59	61	74	67	53	77	61	54	58	50
United States	54	60	58	70	67	61	74	64	57	52	54

Internal Consistency (Cronbach's Alpha) for refined scales in 26 nations

*Note.* Decimal points removed for readability. C = Conscientiousness, H = Honesty/Propriety, A = Agreeableness, R = Resiliency, E = Extraversion, O = Originality/Talent, S = Social Self-Regulation, D = Dynamism. The negative reliability for Big Five Originality in Morocco was because one item "I don't pride myself on being original" had correlations in the wrong direction with all other O items. This was likely due to an Arabic translation that captured the literal meaning, but gave the phrase a different emotional tone and difficulty level ("pride myself" could have been understood as "brag about".)

\* C and A scales are the same for the 30QB6 and 25QB5, so alpha values are not repeated.