Mendelian randomization study of the association between telomere length and risk of cancer and non-neoplastic diseases The Telomeres Mendelian Randomization Collaboration Correspondence: Philip Haycock MRC Integrative Epidemiology Unit University of Bristol Bristol UK philip.haycock@bristol.ac.uk Tel: +44 1173 310 088 2995 words [word limit 3000] 3 figures, 2 tables, 132 references; 7 supplementary figures / 6 supplementary tables 

- **ABSTRACT 349 WORDS**
- 27 **Importance** The causal direction and magnitude of the association between telomere length
- and incidence of cancer and non-neoplastic diseases is uncertain, due to the susceptibility of
- 29 observational studies to confounding and reverse causation.
- 30 **Objective** To conduct a Mendelian randomization study, using germline genetic variants as
- 31 instrumental variables, to appraise the causal relevance of telomere length for risk of cancer
- and non-neoplastic diseases.
- Data Sources Genome-wide association studies (GWAS) published up to January 15 2015.
- 34 Study Selection GWAS of non-communicable diseases that assayed germline genetic
- variation and did not select cohort or control participants on the basis of pre-existing diseases.
- 36 Of 163 GWAS of non-communicable diseases identified, summary data from 103 were
- 37 available.

- 38 **Data Extraction** Summary association statistics for single nucleotide polymorphisms (SNPs)
- that are strongly associated with telomere length in the general population.
- 40 Main Outcomes Odds ratios (ORs) for disease per standard deviation (SD) higher telomere
- 41 length due to germline genetic variation.
- 42 Results Summary data were available for 35 cancers and 48 non-neoplastic diseases,
- corresponding to 420,081 cases (median 2,526 per disease) and 1,093,105 controls (median
- 44 6,789 per disease). Increased telomere length due to germline genetic variation was generally
- 45 associated with increased risk for site-specific cancers. The strongest associations were
- observed for (ORs per 1-SD change in genetically increased telomere length): glioma 5.27
- 47 (3.15-8.81), serous low-malignant-potential ovarian cancer 4.35 (2.39-7.94), lung
- 48 adenocarcinoma 3.19 (2.40-4.22), neuroblastoma 2.98 (1.92-4.62), bladder cancer 2.19 (1.32-
- 49 3.66), melanoma 1.87 (1.55-2.26), testicular cancer 1.76 (1.02-3.04), kidney cancer 1.55

(1.08-2.23) and endometrial cancer 1.31 (1.07-1.61). Associations were stronger for rarer cancers and at tissue sites with lower rates of stem cell division (P<0.05). There was generally little evidence of association between genetically increased telomere length and risk of psychiatric, autoimmune, inflammatory, diabetic and other non-neoplastic diseases, except for coronary heart disease (0.78 [0.67-0.90]), abdominal aortic aneurysm (0.63 [0.49-0.81]), celiac disease (0.42 [0.28-0.61]) and interstitial lung disease (0.09 [0.05-0.15]). Conclusions It is likely that longer telomeres increase risk for several cancers but reduce risk for some non-neoplastic diseases, including cardiovascular diseases. 

#### INTRODUCTION

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At the ends of chromosomes, telomeres are DNA-protein structures that protect the genome from damage, shorten progressively over time in most somatic tissues<sup>1</sup> and are proposed physiological markers of ageing.<sup>2,3</sup> Shorter leukocyte telomeres are correlated with older age, male sex and other known risk factors for non-communicable diseases 4-6 and are generally associated with higher risk for cardiovascular diseases<sup>7,8</sup>, type 2 diabetes<sup>9</sup> and non-vascular non-neoplastic causes of mortality. Whether these associations are causal, however, is unknown. Telomere length has also been implicated in risk of cancer but the direction and magnitude of the association is uncertain and contradictory across observational studies. 10-14 The uncertainty reflects the considerable difficulty of designing observational studies of telomere length and cancer incidence that are robust to reverse causation, confounding and measurement error. The aim of the present report was to conduct a Mendelian randomization study, using germline genetic variants as instrumental variables for telomere length, to help clarify the nature of the association between telomere length and risk of cancer and non-neoplastic diseases. The approach, which mimics the random allocation of individuals to the placebo and intervention arms of a randomized controlled trial, allowed us to: (1) estimate the direction and broad magnitude of the association of telomere length with risk of multiple cancer and non-neoplastic diseases; (2) appraise the evidence for causality in the estimated etiological associations; (3) investigate potential sources of heterogeneity in findings for sitespecific cancers; and (4) compare genetic estimates to findings based on directly measured telomere length in prospective observational studies.

#### **METHODS**

Study design

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The design of our study, illustrated in Figure S1, had three key components: 1) the 98 99 identification of genetic variants to serve as instruments for telomere length; 2) the 100 acquisition of summary data for the genetic instruments from genome wide association 101 studies (GWASs) of diseases and risk factors for non-communicable diseases; and 3) the 102 classification of diseases and risk factors into primary or secondary outcomes based on a priori statistical power. As a first step, we searched the GWAS catalog <sup>15,16</sup> on the 15 January 103 104 2015, to identify single nucleotide polymorphisms (SNPs) associated with telomere length. 105 To supplement the list with additional potential instruments, we also searched the original study reports curated by the GWAS catalog (using a P-value threshold of 5x10<sup>-8</sup>). 17-25 We 106 acquired summary data for all SNPs identified by our search from a meta-analysis of GWASs 107 of telomere length, involving 9,190 participants of European ancestry. 18 108 The second key component of our design strategy involved the acquisition of summary data, 109 110 corresponding to the selected genetic instruments for telomere length, from GWASs of noncommunicable diseases and risk factors (Fig. S1). As part of this step, we invited principal 111 investigators of non-communicable disease studies curated by the GWAS catalog 15,26 to share 112 summary data for our study (see Fig. S1 for further details). We also downloaded summary 113 114 data for diseases and risk factors from publically available sources, including study-specific websites, dbGAP, ImmunoBase and the GWAS catalog (Fig. S1). 115 The third key component of our design strategy was the classification of diseases and risk 116 117 factors into either primary or secondary outcomes, which we defined on the basis of a priori statistical power to detect associations with telomere length. Primary outcomes were defined 118

as diseases with sufficient cases and controls for >50% statistical power and secondary outcomes defined as diseases with <50% statistical power to detect odds ratios  $\ge 2.0$  per standard deviation (SD) change in genetically increased telomere length (alpha assumed to be 0.01). All risk factors were defined as secondary outcomes. Risk factors with <50% statistical power were excluded.

Further details on our design strategy can be found in the supplement.

Comparison with prospective observational studies

We searched PubMed for prospective observational studies of the association between telomere length and disease (see Tables S3 and S4 for details of the search strategy and inclusion criteria). Study-specific relative risks for disease per unit change or quantile comparison of telomere length were transformed to a SD scale using previously described methods. Hazard ratios, risk ratios and odds ratios were assumed to approximate the same measure of relative risk. Where multiple independent studies of the same disease were identified, these were combined by fixed effects meta-analysis, unless there was strong evidence of between-study heterogeneity ( $P_{Cochran's Q}$ <0.001), in which case they were kept separate.

Statistical analysis

We combined summary data across SNPs into a single instrument, using maximum likelihood to estimate the slope of the relationship between  $\beta_{GD}$  and  $\beta_{GP}$  and a variance-covariance matrix to make allowance for linkage disequilibrium between SNPs, <sup>28</sup> where  $\beta_{GD}$  is the change in disease log odds or risk factor levels per copy of the effect allele and  $\beta_{GP}$  is the SD change in telomere length per copy of the effect allele (see supplementary methods

for technical details). The slope from this approach can be interpreted as the log odds ratio for binary outcomes, or the unit change for continuous risk factors, per SD change in genetically increased telomere length. P-values for heterogeneity amongst SNPs, in the estimated associations of genetically increased telomere length with disease and risk factors, were estimated by likelihood ratio tests.<sup>28</sup> Associations between genetically increased telomere length and continuous risk factors were transformed into SD units. For five secondary disease outcomes where only a single SNP was available for analysis, we estimated associations using the Wald ratio:  $\beta_{GD}/\beta_{GP}$ , with standard errors approximated by the delta method.<sup>29</sup> Inference of causality in the estimated etiological associations between telomere length and disease depends on satisfaction of Mendelian randomization assumptions (Fig. S7; see Table S6 for a glossary of terms). 30,31 The assumptions are: 1) the selected SNPs are associated with telomere length; 2) the selected SNPs are not associated with confounders; and 3) the selected SNPs are associated with disease exclusively through their effect on telomere length. If these assumptions are satisfied, the selected SNPs are valid instrumental variables and their association with disease can be interpreted as a causal effect of telomere length. We modeled the impact of violations of these assumptions through two sets of sensitivity analyses: a weighted median function<sup>32</sup> and MR-Egger regression<sup>30</sup> (see supplementary methods for technical details). We restricted our sensitivity analyses to diseases showing the strongest evidence of association with genetically increased telomere length (defined as  $P_{Bonferroni} \leq 0.05$ ).

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We used meta-regression to appraise potential sources of heterogeneity in our findings for cancer. The association of genetically increased telomere length with the log odds of cancer was regressed on cancer incidence, survival time and median age-at-diagnosis, downloaded from the National Cancer Institute Surveillance, Epidemiology, and End Results (SEER)

Program,<sup>33</sup> and tissue-specific rates of stem cell division from Tomasetti and Vogelstein.<sup>34</sup> As the downloaded cancer characteristics from SEER correspond to the United States population, 77% of which was of white ancestry in 2015<sup>35</sup>, the meta-regression analyses excluded genetic studies conducted in East Asian populations.

All analyses were performed in R version 3.1.2<sup>36</sup> and Stata release 13.1 (StataCorp, College Station, TX). P-values were two-sided and evidence of association was declared at P<0.05. Where indicated, Bonferroni corrections were used to make allowance for multiple testing, although this is likely to be overly conservative given the non-independence of many of the outcomes tested.

# **RESULTS**

We selected 16 SNPs as instruments for telomere length (Fig. S1 & Table 1). The selected SNPs correspond to 10 independent genomic regions that collectively account for 2-3% of the variance in leukocyte telomere length, which is equivalent to an F statistic of ~18. This indicates that the genetic instrument, constructed from these 10 independent genomic regions, is strongly associated with telomere length (details in supplementary discussion). Summary data for the genetic instruments were available for 83 non-communicable diseases, corresponding to 420,081 cases (median 2,526 per disease) and 1,093,105 controls (median 6,789 per disease), and 44 risk factors (Fig. S1, Table 2 and Table S1). The median number of SNPs available across diseases was 11 (min=1, max=13) and across risk factors was 12 (min=11, max=13). Of the 83 diseases, 56 were classified as primary outcomes and 27 as secondary outcomes (Table 2, Fig. S1 and Table S1). For 9 of the 83 non-communicable diseases, additional summary data were available from 10 independent studies for replication

analyses, corresponding to 40,465 cases (median 1,416 per disease) and 52,306 controls (median 3,537 per disease) (Table S1). 195 196 The results from primary analyses of non-communicable diseases are presented in Figure 1; 197 results from secondary analyses of risk factors and diseases with low a priori power are presented in the supplement (Fig. S2, S5 and S6). Genetically increased telomere length was 198 199 associated with higher odds of disease for 9 of 22 primary cancers (P<0.05), including (odds 200 ratio [95% confidence interval]): glioma (5.27 [3.15-8.81]), endometrial cancer (1.31 [1.07-1.61]), kidney cancer (1.55 [1.08-2.23]), testicular germ cell cancer (1.76 [1.02-3.04]), 201 202 melanoma (1.87 [1.55-2.26]), bladder cancer (2.19 [1.32-3.66]), neuroblastoma (2.98 [1.92-203 4.62]), lung adenocarcinoma (3.19 [2.40-4.22]) and serous low-malignancy-potential (LMP) 204 ovarian cancer (4.35 [2.39-7.94]) (Fig. 1). The associations were, however, highly variable 205 across cancer types, varying from an odds ratio of 0.86 (0.50-1.48) for head and neck cancer 206 to 5.27 (3.15-8.81) for glioma. Substantial variability was also observed within tissue sites. 207 For example, the odds ratio for lung adenocarcinoma was 3.19 (2.40-4.22) compared to 1.07 208 (0.82-1.39) for squamous cell lung cancer. For serous LMP ovarian cancer the odds ratio was 209 4.35 (2.39-7.94) compared to odds ratios of 1.21 (0.87-1.68) for endometrioid ovarian cancer, 210 1.12 (0.94-1.34) for serous invasive ovarian cancer, 1.04 (0.66-1.63) for clear cell ovarian 211 cancer and 1.04 (0.73-1.47) for mucinous ovarian cancer. The strongest evidence of 212 association was observed for glioma, lung adenocarcinoma, neuroblastoma and serous LMP ovarian cancer (P<sub>Bonferroni</sub><0.05). Results for glioma and bladder cancer showed evidence for 213 replication in independent datasets (independent datasets were not available for other 214 cancers) (Fig. S3). 215 216 Genetically increased telomere length was associated with reduced odds of disease for 6 of 32 217 primary non-neoplastic diseases (P<0.05), including coronary heart disease (0.78 [0.67-0.9]), 218 abdominal aortic aneurysm (0.63 [0.49-0.81]), Alzheimer's disease (0.84 [0.71-0.98]), celiac

disease (0.42 [0.28-0.61]), interstitial lung disease (0.09 [0.05-0.15]) and type 1 diabetes (0.71 [0.51-0.98]) (P<0.05) (Figure 1). The strongest evidence of association was observed for coronary heart disease (P<sub>Bonferroni</sub>=0.05) and abdominal aortic aneurysm, celiac disease and interstitial lung disease (P<sub>Bonferroni</sub><0.05). The associations with coronary heart disease and interstitial lung disease showed evidence for replication in independent datasets (Fig. S3). Our genetic findings were generally similar in direction and magnitude to estimates based on observational prospective studies of leukocyte telomere length and disease (Figure 3). Our genetic estimates for lung adenocarcinoma, melanoma, kidney cancer and glioma, were, however, stronger in comparison to observational estimates. In sensitivity analyses, we appraised the potential impact of confounding by pleiotropic pathways on our results. Associations estimated by the weighted median and MR-Egger were broadly similar to the main results for glioma, lung adenocarcinoma, serous LMP ovarian cancer, neuroblastoma, abdominal aortic aneurysm, coronary heart disease and interstitial lung disease (Fig. S4). In the second set of sensitivity analyses, implemented by MR-Egger regression, we found little evidence for the presence of pleiotropy (P<sub>intercept</sub>≥0.27) (Fig. S4). The MR-Egger analyses were, however, generally underpowered, as reflected by the wide confidence intervals in the estimated odds ratios. In meta-regression analyses, we observed that genetically increased telomere length tended to

be more strongly associated with rarer cancers (P=0.02) and cancers at tissue-sites with lower

rates of stem cell division (P=0.02) (Figure 2). The associations showed little evidence of

varying by percentage survival five years after diagnosis or median age-at-diagnosis ( $P \ge 37$ ).

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#### **DISCUSSION**

In this report we show that genetically increased telomere length is associated with increased risk of several cancers and with reduced risk of some non-neoplastic diseases. Given the random distribution of genotypes in the general population with respect to lifestyle and other environmental factors, as well as the fixed nature of germline genotypes, these results should be less susceptible to confounding and reverse causation in comparison to observational studies. Our results are therefore compatible with causality. On the other hand, our results could reflect violations of Mendelian randomization assumptions, such as confounding by pleiotropy, population stratification or ancestry. Although we cannot entirely rule out this possibility, the majority of our results persisted in sensitivity analyses that made allowance for violations of Mendelian randomization assumptions. Confounding by population stratification or ancestry is also unlikely, given the adjustments made for ancestry in the disease GWASs (see supplementary discussion).

# Comparison with previous studies

Our findings for cancer are generally contradictory to those based on retrospective studies, which tend to report increased risk for cancer in individuals with shorter telomeres. 11,12,39–42

The contradictory findings may reflect reverse causation in the retrospective studies, whereby shorter telomeres arise as a result of disease, or of confounding effects, e.g. due to cases being slightly older than controls even in age-matched analyses. Our findings for cancer are generally more consistent with those based on prospective observational studies, which tend to report weak or null associations of longer leukocyte telomeres with overall and site-specific risk of cancer, 10–13,41,43–62 with some exceptions. 63 Our results are also similar to previously reported Mendelian randomization studies of telomere length and risk of

melanoma, lung cancer, chronic lymphocytic leukemia and glioma. 64-67 The shape of the association with cancer may not, however, be linear over the entire telomere length distribution. For example, individuals with dyskeratosis congenita, a disease caused by germline loss-of-function mutations in the telomerase component genes *TERC* and *TERT*, have chronically short telomeres and are at increased risk of some cancers, particularly acute myeloid leukemia and squamous cell carcinomas arising at sites of leukoplakia, 68,69 presumably due to increased susceptibility to genome instability and chromosomal end-to-end fusions. 70 Our results should therefore be interpreted as reflecting the average association at the population level and may not be generalizable to the extreme ends of the telomere length distribution.

# Mechanisms of association

Our cancer findings are compatible with known biology. <sup>70</sup> By limiting the proliferative potential of cells, telomere shortening may serve as a tumour suppressor; and individuals with longer telomeres may be more likely to acquire somatic mutations owing to increased proliferative potential. <sup>70</sup> Rates of cell division are, however, highly variable amongst tissues <sup>34</sup> and thus the relative gain in cell proliferative potential, conferred by having longer telomeres, may also be highly variable across tissues. This could explain the ~6-fold variation in odds ratios observed across cancer types in the present study, as well as the tendency of our results to be stronger at tissue sites with lower rates of stem cell division. For example, the association was strongest for glioma (OR=5.27) and comparatively weak for colorectal cancer (OR=1.09) and the rates of stem cell division in the tissues giving rise to these cancers differ by several orders of magnitude. In neural stem cells, which give rise to gliomas, the number of divisions is ~270 million and for colorectal stem cells is ~1.2 trillion over the average lifetime of an individual. <sup>34</sup> The observation that genetically increased telomere

length was more strongly associated with rarer cancers potentially reflects the same mechanism, since rarer cancers also tend to show lower rates of stem cell division.<sup>34</sup> For example, the incidence of glioma is 0.4 and for colorectal cancer is 42.4 per 100,000 per year in the United States.<sup>33</sup>

The inverse associations observed for some non-neoplastic diseases may reflect the impact of telomere shortening on tissue degeneration and an evolutionary trade-off for greater resistance to cancer at the cost of greater susceptibility to degenerative diseases, particularly cardiovascular diseases. 71,72

# Study limitations

Our study is subject to some limitations, in addition to the Mendelian randomization assumptions already considered above. First, our method assumes that the magnitude of the association between SNPs and telomere length is consistent across tissues. Second, our study assumed a linear shape of association between telomere length and disease risk, whereas the shape could be "J" or "U" shaped. 44,57,68 Third, our results assume that the samples used to define the genetic instrument for telomere length and the various samples used to estimate the SNP-disease associations are representative of the same general population, practically defined as being of similar ethnicity, age and sex distribution. This assumption would, for example, not apply in the case of the SNP-disease associations derived from East Asian or pediatric populations. Generally speaking, violation of the aforementioned assumptions could bias the magnitude of the association between genetically increased telomere length and disease; but would be unlikely to increase the likelihood of false positives (i.e. incorrectly inferring an association when none exists). Our results should therefore remain informative for the direction and broad magnitude of the average association at the population level, even

in the presence of such violations. Fourth, we cannot rule out chance in explaining some of the weaker findings. Fifth, our results may not be fully representative of non-communicable diseases (since not all studies shared data and our analyses were underpowered for the secondary disease outcomes). The diseases represented in our primary analyses probably account for >60% of all causes of death in American adults.<sup>75</sup>

#### Clinical relevance of findings

Our findings suggest that potential clinical applications of telomere length, e.g. as a tool for risk prediction or as an intervention target for disease prevention, may have to consider a trade-off in risk between cancer and non-neoplastic diseases. For example, a number of companies have been established that offer telomere length measurement services to the public (via a requesting physician), under the claim that shorter telomeres are a general indicator of poorer health status and older biological age and that such information can be used to motivate healthy lifestyle choices in individuals. However, the conflicting direction of association between telomere length and risk of cancer and non-neoplastic diseases, indicated by our findings, suggests that such services to the general public may be premature.

#### Conclusion

It is likely that longer telomeres increase risk for several cancers but reduce risk for some non-neoplastic diseases, including cardiovascular diseases. Further research is required to resolve whether telomere length is a useful predictor of risk that can help guide therapeutic interventions, to clarify the shape of any dose-response relationships and to characterise the nature of the association in population subgroups.

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Table 1. Single nucleotide polymorphisms associated with telomere length

											No.	Sample	Discovery	% variance	
SNPs	Chr	Pos	Gene	EA	OA	EAF*	Beta*	SE*	P-value*	Phet*	studies*	size*	p-value	explained	Discovery study
rs11125529	2	54248729	ACYP2	A	C	0.16	0.065	0.012	0.000606	0.313	6	9177	8.00E-10	0.080	$Codd^{21}$
rs6772228	3	58390292	PXK	T	A	0.87	0.041	0.014	0.049721	0.77	6	8630	3.91E-10	0.200	Pooley <sup>17</sup>
rs12696304	3	169763483	TERC	C	G	0.74	0.090	0.011	5.41E-08	0.651	6	9012	4.00E-14	0.319	$Codd^{22}$
rs10936599	3	169774313	TERC	C	T	0.76	0.100	0.011	1.76E-09	0.087	6	9190	3.00E-31	0.319	$Codd^{21}$
rs1317082	3	169779797	TERC	A	G	0.71	0.097	0.011	4.57E-09	0.029	6	9176	1.00E-08	0.319	Mangino <sup>18</sup>
rs10936601	3	169810661	TERC	C	T	0.74	0.087	0.011	8.64E-08	0.433	6	9150	4.00E-15	0.319	Pooley <sup>17</sup>
rs7675998	4	163086668	NAF1	G	A	0.80	0.048	0.012	0.008912	0.077	6	9161	4.35E-16	0.190	$Codd^{21}$
rs2736100	5	1286401	TERT	C	A	0.52	0.085	0.013	2.14E-05	0.54	4	5756	4.38E-19	0.310	$Codd^{21}$
rs9419958	10	103916188	OBFC1	T	C	0.13	0.129	0.013	5.26E-11	0.028	6	9190	9.00E-11	0.171	Mangino <sup>18</sup>
rs9420907	10	103916707	OBFC1	C	A	0.14	0.142	0.014	1.14E-11	0.181	6	9190	7.00E-11	0.171	$Codd^{21}$
rs4387287	10	103918139	OBFC1	A	C	0.14	0.120	0.013	1.40E-09	0.044	6	8541	2.00E-11	0.171	Levy <sup>25</sup>
rs3027234	17	8232774	CTC1	C	T	0.83	0.103	0.012	2.75E-08	0.266	6	9108	2.00E-08	0.292	Mangino <sup>18</sup>
rs8105767	19	22032639	ZNF208	G	A	0.25	0.064	0.011	0.000169	0.412	6	9096	1.11E-09	0.090	$Codd^{21}$
rs412658	19	22176638	ZNF676	T	C	0.35	0.086	0.010	1.83E-08	0.568	6	9156	1.00E-08	0.484	Mangino <sup>18</sup>
rs6028466	20	39500359	DHX35	A	G	0.17	0.058	0.013	0.003972	0.533	6	9190	2.57E-08†	0.041	Mangino <sup>18</sup> & G
rs755017	20	63790269	ZBTB46	G	A	0.17	0.019	0.0129	0.339611	0.757	5	8026	6.71E-09	0.090	$Codd^{21}$

<sup>\*</sup>Summary data from Mangino et al<sup>18</sup>; Chr, chromosome; pos, base-pair position (GRCh38.p3); EA, effect allele, OA, other allele, Beta, standard deviation change in telomere length per copy of the effect allele; SE, standard error; EAF - effect allele frequency; Phet - p value for between-study heterogeneity in association between SNP and telomere length; †from a meta-analysis of Mangino 18 and Gu<sup>20</sup> performed in the present study.

Table 2. Study characteristics for primary non-communicable diseases

	No.	No.	No.	Statistical	D.	Ct. 1. / E: 1
	cases	controls	SNPs	power	Pop.	Study / First author
Cancer	1.601	1010	10	0.62	ELID	NIDGG76
Bladder cancer	1601	1819	10	0.62	EUR	NBCS <sup>76</sup>
Breast cancer	48155	43612	13	1.00	EUR	BCAC <sup>17,77</sup> BCAC <sup>17,77</sup>
Estrogen receptor –ve	7465	42175	13	1.00	EUR	BCAC <sup>17,77</sup>
Estrogen receptor +ve	27074	41749	13	1.00	EUR	
Colorectal cancer	14537	16922	9	1.00	EUR	CORECT/GECCO <sup>64,78</sup> ECAC <sup>79,80</sup>
Endometrial cancer	6608	37925	12	1.00	EUR	Abnet <sup>81</sup>
Esophageal SCC	1942	2111	11	0.64	EA	Wrensch <sup>82</sup> & Walsh <sup>66</sup>
Glioma	1130	6300	12	0.72	EUR	
Head & neck cancer	2082	3477	12	1.00	EUR	McKay et al <sup>83</sup>
Kidney cancer	2461	5081	12	0.99	EUR	KIDRISK <sup>84</sup>
Lung cancer	11348	15861	13	1.00	EUR	ILCCO <sup>85</sup> ILCCO <sup>85</sup>
Adenocarcinoma	3442	14894	13	1.00	EUR	
Squamous cell carcinoma Skin cancer	3275	15038	13	1.00	EUR	ILCCO <sup>85</sup>
Melanoma	12814	23203	13	1.00	EUR	$MC^{86}$
Basal cell carcinoma	3361	11518	13	1.00	EUR	NHS/HPFS <sup>87</sup>
Neuroblastoma	2101	4202	12	0.87	EUR	Diskin <sup>88</sup>
Ovarian cancer	15397	30816	13	1.00	EUR	OCAC <sup>17,89</sup>
Clear cell	1016	30816	13	0.76	EUR	OCAC <sup>17,89</sup>
Endometriod	2154	30816	13	0.98	EUR	OCAC <sup>17,89</sup>
Mucinous	1643	30816	13	0.94	EUR	OCAC <sup>17,89</sup>
Serous invasive	9608	30816	13	1.00	EUR	OCAC <sup>17,89</sup>
Serous LMP	972	30816	13	0.73	EUR	OCAC <sup>17,89</sup>
Pancreatic cancer	5105	8739	12	1.00	EUR	PanScan (incl. EPIC) <sup>90</sup>
Prostate cancer	22297	22323	11	1.00	EUR	PRACTICAL <sup>91,92</sup>
Testicular germ cell cancer	986	4946	11	0.52	EUR	Turnbull <sup>93</sup> & Rapley <sup>94</sup>
Autoimmune/inflammatory dis				***-		
Alopecia areata	2332	5233	7	0.60	EUR	Betz <sup>95</sup>
Atopic dermatitis	10788	30047	13	1.00	EUR	EAGLE <sup>96</sup>
Celiac disease	4533	10750	3	0.82	EUR	Dubois <sup>97</sup>
Inflammatory bowel disease						
Crohn's disease	5956	14927	11	1.00	EUR	IIBDGC <sup>98</sup>
Ulcerative colitis	6968	20464	12	1.00	EUR	IIBDGC <sup>98</sup>
Juvenile idiopathic arthritis	1866	14786	11	0.87	EUR	Thompson <sup>99</sup> †
Multiple sclerosis	14498	24091	3	1.00	EUR	IMSGC <sup>100</sup>
Aggressive periodontitis	888	6789	13	0.63	EUR	Schaefer <sup>101</sup>
Rheumatoid arthritis	5538	20163	11	1.00	EUR	Stahl <sup>102</sup>
Cardiovascular diseases	3336	20103	11	1.00	LUK	Stani
Abdominal aortic aneurysm	4972	99858	13	1.00	EUR	$AC^{103-108}$
Coronary heart disease	22233	64762	13	1.00	EUR	CARDIoGRAM <sup>109</sup>
Heart failure	2526	20926	13	0.99	EUR	CHARGE-HF <sup>110</sup>
Hemorrhagic stroke	2963	5503	12	0.96	EUR	METASTROKE/ISGC <sup>111</sup>
Ischemic stroke	12389	62004	13	1.00	EUR	METASTROKE/ISGC <sup>112</sup> ,
large vessel disease	2167	62004	13	0.99	EUR	METASTROKE/ISGC <sup>112</sup> ,
small vessel disease	1894	62004	13	0.99	EUR	METASTROKE/ISGC <sup>112</sup>
cardioembolic	2365	62004	13	0.97	EUR	METASTROKE/ISGC <sup>112</sup>
Sudden cardiac arrest	3954	21200	13	1.00	EUR	Unpublished
Diabetes	3334	21200	13	1.00	LUK	Onpublished
Type 1 diabetes	7514	9045	6	0.95	EUR	T1DBase <sup>114115</sup>
Type 2 diabetes	10415	53655	0 11	1.00	EUR	DIAGRAM <sup>116</sup>
Eye disease	10+13	55055	11	1.00	LUK	DIAUKAM

AMD	7473	51177	13	1.00	EUR	AMD Gene <sup>117</sup>			
Retinopathy	1122	18289	12	0.75	EUR	Jensen <sup>118</sup>			
Lung diseases									
Asthma	13034	20638	4	1.00	EUR	Ferreira/GABRIEL <sup>119,120</sup>			
COPD	2812	2534	12	0.85	EUR	COPDGene <sup>121</sup>			
Interstitial lung disease	1616	4683	9	0.60	EUR	Fingerlin <sup>122</sup>			
Neurological / psychiatric diseases									
ALS	6100	7125	12	1.00	EUR	SLAGEN/ALSGEN <sup>123</sup>			
Alzheimer's disease	17008	37154	12	1.00	EUR	IGAP <sup>124</sup>			
Anorexia nervosa	2907	14860	9	0.93	EUR	GCAN <sup>125</sup>			
Autism	4949	5314	7	0.82	EUR	$PGC^{126}$			
Bipolar disorder	7481	9250	9	1.00	EUR	$PGC^{127}$			
Major depressive disorder	9240	9519	8	0.99	EUR	$PGC^{128}$			
Schizophrenia	35476	46839	12	1.00	EUR	$PGC^{129}$			
Tourette syndrome	1177	4955	13	0.74	EUR	TICG/TSAICG <sup>130</sup>			
Other									
Chronic kidney disease	5807	56430	13	1.00	EUR	CKDGen <sup>131</sup>			
Endometriosis	4604	9393	11	1.00	Mix	Nyholt <sup>132</sup>			

Study acronyms: AC, the aneurysm consortium; ALSGEN, the International Consortium on Amyotrophic Lateral Sclerosis Genetics; AMD Gene, Age-related Macular Degeneration Gene Consortium; BCAC, Breast Cancer Association Consortium; CARDIoGRAM, Coronary ARtery DIsease Genome wide Replication and Meta-analysis; CHARGE-HF, Cohorts for Heart and Aging Research in Genomic Epidemiology Consortium - Heart Failure Working Group; COPDGene, The Genetic Epidemiology of Chronic Obstructive Pulmonary Disease; CKDGen, Chronic Kidney Disease Genetics consortium; CORECT, ColoRectal Transdisciplinary Study; DIAGRAM, DIAbetes Genetics Replication And Meta-analysis; EAGLE, EArly Genetics & Lifecourse Epidemiology Eczema Consortium (excluding 23andMe); ECAC, Endometrial Cancer Association Consortium; EPIC, European Prospective Investigation into Cancer and Nutrition study; GABRIEL, Multidisciplinary Study to Identify the Genetic and Environmental Causes of Asthma in the European Community; GCAN, Genetic Consortium for Anorexia Nervosa; GECCO, Genetics and Epidemiology of Colorectal Cancer Consortium; IGAP, International Genomics of Alzheimer's Project; HPFS, Health Professionals Follow-Up Study; ILCCO, International Lung Cancer Consortium; IMSGC, International Multiple Sclerosis Genetic Consortium; IIBDGC, International Inflammatory Bowel Disease Genetics Consortium; KIDRISK, Kidney cancer consortium; MC, the melanoma meta-analysis consortium; METASTROKE/ISGC, METASTROKE project of the International Stroke Genetics Consortium; NBCS, Nijmegen Bladder Cancer Study; NHS, Nurses' Health Study; OCAC, Ovarian Cancer Association Consortium; PanScan, Pancreatic Cancer Cohort Consortium; PGC, Psychiatric Genomics Consortium; PRACTICAL, Prostate Cancer Association Group to Investigate Cancer Associated Alterations in the Genome; SLAGEN, Italian Consortium for the Genetics of Ayotrophic Lateral Sclerosis; T1DBase, type 1 diabetes database; TICG (Tourette International Collaborative-Genetics); TSAICG (Tourette Syndrome Association International Consortium for Genetics);. Abbreviations: ALS, amyotrophic lateral sclerosis; AMD, age-related macular degeneration; COPD, chronic obstructive pulmonary disease; EUR, European; EA, East Asian; LMP, low malignant potential; No., number; Pop., population; SCC, squamous cell carcinoma; SNP, single nucleotide polymorphism; -ve, negative; +ve, positive; †plus previously unpublished data.

# Figure 1. The association between genetically increased telomere length and odds of primary non-communicable diseases

# Legend to Figure 1

\*P value for association between genetically increased telomere length and disease from maximum likelihood; the effect estimate for heart failure is a hazard ratio (all others are odds ratios);  $P_{het}$ , P-value for heterogeneity amongst SNPs within the instrument; COPD, chronic obstructive pulmonary disease; SNP, single nucleotide polymorphism; CI, confidence interval; LMP, low malignancy potential; ER, estrogen receptor; -VE, negative; +VE, positive.

# **Figure 2.** The association between genetically increased telomere length and odds of cancer as a function of selected characteristics

#### Legend to Figure 2

The plotted data show how the strength of the relationship between genetically increased telomere length and cancer varies by the selected characteristic. The R² statistic indicates how much of the variation between cancers can be explained by the selected characteristic. P-values are from meta-regression models. Circle sizes are proportional to the inverse of the variance of the log odds ratio. The hashed line indicates the null of no association between telomere length and cancer (i.e. an odds ratio of 1). Data for percentage survival 5 years after diagnosis, cancer incidence and median age-at-diagnosis was downloaded from the Surveillance, Epidemiology, and End Results Program.<sup>33</sup> Data for average lifetime number of stem cell divisions was downloaded from Tomasetti and Vogelstein.<sup>34</sup> Not all cancers had information available for the selected characteristics (hence the number of cancers varies across the subplots). Information was available for 9 cancers for tissue-specific rates of stem cell division, 13 cancers for percentage surviving 5 years post-diagnosis, 17 cancers for cancer incidence and 13 cancers for median age-at-diagnosis. SD, standard deviation; OR, Odds ratio.

# **Figure 3.** Comparison of genetic and prospective observational studies<sup>†</sup> of the association between telomere length and disease

# Legend to Figure 3

\*from fixed-effects meta-analysis of independent observational studies described in Table S3; †search strategy and characteristics for observational studies are described in Tables S3 and S4; ‡CCHS and CGPS; +PLCO, ATBC & SWHS (acronyms explained in Table S3); CL confidence interval

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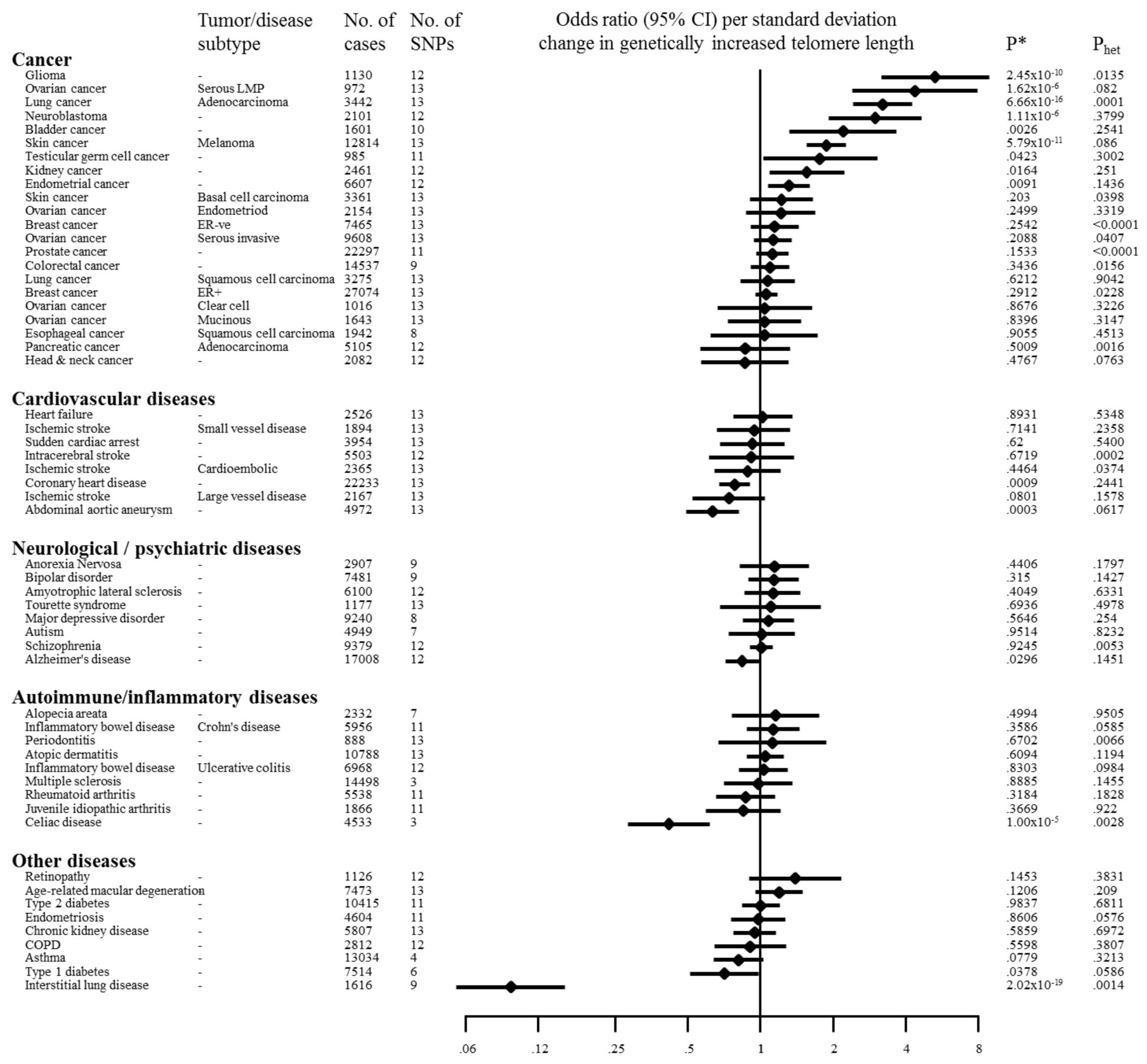
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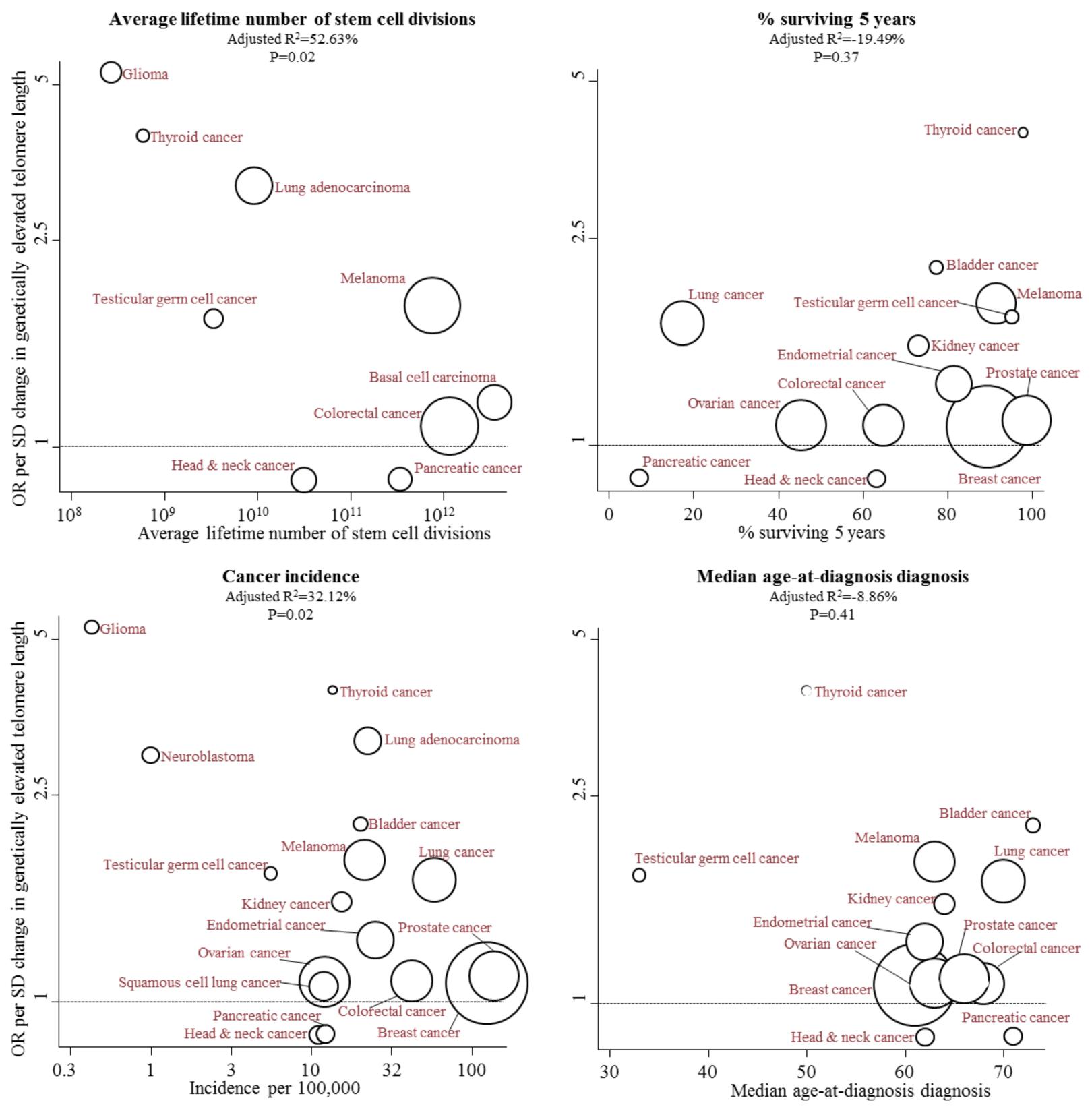
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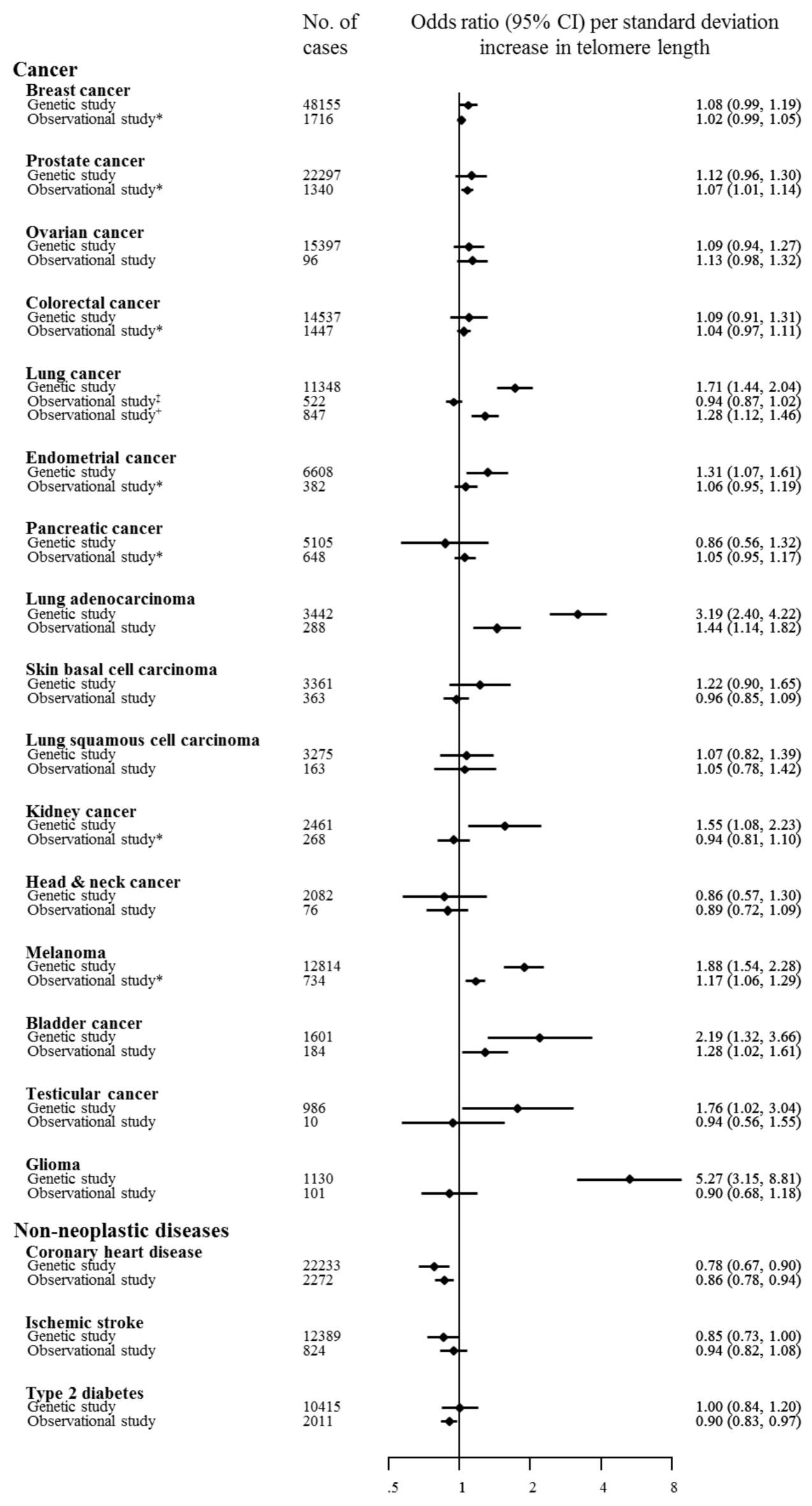
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1	Supplementary ma	aterial
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3	and non-neoplastic	e diseases
4	T1 T1 V	
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#### SUPPLEMENTARY METHODS

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# Additional details on the design strategy

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Identification of genetic instruments for telomere length

To identify genetic variants to serve as instruments for telomere length, we searched the genomewide association study (GWAS) catalog<sup>1,2</sup> on the 15 January 2015, to identify reported single nucleotide polymorphisms (SNPs) associated with telomere length. To supplement the list with additional potential instruments, we also searched the original study reports curated by the GWAS catalog. 3-11 We included all 'telomere length' SNPs in the GWAS catalog as potential proxies, regardless of their reported P-value, but used a P-value threshold of  $<5x10^{-8}$  (the conventional threshold for declaring association in GWAS) for SNPs identified from original study reports (if these were not already curated by the GWAS catalog). We acquired summary data for all SNPs identified by the above strategy from a meta-analysis of six GWASs of leukocyte telomere length, conducted in 9,190 participants of European ancestry.<sup>4</sup> Telomere length in the six studies was measured by Southern blotting. GWAS analyses in the 6 studies were adjusted for age, sex, body mass index and smoking history. The genomic control inflation factor ( $\lambda_{GC}$ ) ranged from 0.995 to 1.076 across the six studies, indicating little evidence for confounding by population stratification.<sup>4</sup> The following summary data were acquired for each SNP from each of the six studies: the regression coefficient (beta) and its standard error, where the beta reflects the change in telomere length (in base pair units) per copy of the effect allele; the effect allele; the non-effect allele; and effect allele frequency. We combined the effect estimates from the six separate studies by fixed effects meta-analysis. We then excluded SNPs if they lacked strong evidence of association with telomere length. We defined strong evidence of association as a P value <5x10<sup>-8</sup> in: i) the discovery stage of at least one published GWAS of telomere length<sup>3–10</sup> or ii) a meta-analysis of summary data

from Mangino et al<sup>4</sup> and other GWASs of telomere length,  $^{3,5-10}$  with any overlapping studies excluded from Mangino et al.<sup>4</sup> We also excluded SNPs with a minor allele frequency <0.05 or showing strong evidence of between-study heterogeneity in associations with telomere length (P $\leq$ 0.001).

Acquisition of summary data from disease and risk factor studies

We extracted the following summary data for each genetic instrument for telomere length from GWASs of diseases and risk factors: the regression coefficient (beta) and its standard error, the effect allele, the non-effect allele and effect allele frequency. For binary traits, the beta corresponded to the log odds ratio per copy of the effect allele. For quantitative traits, the beta corresponded to the unit change in the trait per copy of the effect allele. We harmonized the summary data for diseases and risk factors so that the effect allele reflected the allele associated with longer telomeres. When SNPs were palindromic, i.e. A/T or G/C, we used information on allele frequency to resolve strand ambiguity. We also requested the following metrics of SNP genotype quality: P-values for Hardy-Weinberg equilibrium (HWE), imputation quality scores and P-values for between-study heterogeneity. We also estimated the percentage overlap in participants amongst the telomere length and disease and risk factor GWASs. When reported, statistics on between-study heterogeneity, Hardy-Weinberg equilibrium and imputation quality were used to exclude low quality SNPs from disease and risk factor studies, using the following criteria: strong evidence of between-study heterogeneity in the SNP-phenotype association ( $P \le 0.001$ ), Hardy-Weinberg disequilibrium ( $P \le 0.001$ ) or imputation quality metric (info or  $P \ge 0.90$ ).

Power calculations

Power calculations for disease outcomes were implemented using the method described by Burgess<sup>12</sup> and assumed an odds ratio of  $\geq$ 2.0 per standard deviation higher telomere length and an alpha of 0.01. Power calculations for risk factors for non-communicable diseases were similar,

except that a  $\geq$ 0.5 standard deviation change in quantitative risk factors and an odds ratio of  $\geq$ 1.5 for binary risk factors was assumed for each standard deviation change in telomere length. When more than one study was available for the same outcome trait, priority was given to the study with the higher statistical power. Power calculations took into account the variance explained in telomere length by each SNP, inferred from published reports,<sup>3–10</sup> and the sample size available for each outcome.

#### Estimating the association between genetically increased telomere length and outcome traits

We employed three general approaches for estimating the association between genetically increased telomere length and outcome traits. Our main results are based on a likelihood-approach.<sup>13</sup> Sensitivity analyses were based on two approaches: the weighted median<sup>14</sup> and MR-Egger regression.<sup>15</sup> The technical details of these approaches are described below.

Prior to calculating the associations of genetically increased telomere length with diseases and risk factors, we estimated the pairwise r<sup>2</sup> for all telomere-associated SNPs residing on the same chromosome using PLINK<sup>16</sup> and 1000 Genomes phase 3 data for European samples.<sup>17</sup> SNPs residing on separate chromosomes or separated by more than 50 megabases on the same chromosome were assumed to be in linkage equilibrium. The genetic instruments for telomere length were pruned so that no SNP pair had an r<sup>2</sup>>0.9 (strong linkage disequilibrium), using the 'indep' command in PLINK.<sup>16</sup> The base pair position and chromosome id for each SNP, in GCRCh38 format, was extracted from Ensembl through the R biomart package.<sup>18–20</sup> Linkage disequilibrium between the remaining SNPs was taken into account using a variance-covariance matrix (described below). For analyses in which SNP-disease associations were derived from East Asian populations, genetic instruments were further pruned so that no SNP pair had an r<sup>2</sup>>0.1 (because the variance-covariance matrix used to model the correlation between SNPs was based on a European population).

Likelihood approach

We combined summary data across SNPs into a single instrument, using maximum likelihood to estimate the slope of the relationship between  $\beta_{GD}$  and  $\beta_{GP}$  and a variance-covariance matrix to make allowance for linkage disequilibrium between SNPs, where  $\beta_{GD}$  is the change in the outcome trait per copy of the effect allele and  $\beta_{GP}$  is the standard deviation change in telomere length per copy of the effect allele.<sup>13</sup> The standard deviation of telomere length corresponds to approximately 650 base pairs.<sup>4</sup> The variance-covariance matrix was estimated using 1000 Genomes phase 3 data for Europeans.<sup>13</sup> The model that is fitted is:

$$\begin{pmatrix} \boldsymbol{\beta_{GP}} \\ \boldsymbol{\beta_{GD}} \end{pmatrix} \sim N_{2K} \begin{pmatrix} \boldsymbol{\xi} \\ \boldsymbol{\beta_{IV}\boldsymbol{\xi}} \end{pmatrix}, \begin{pmatrix} \boldsymbol{\Sigma_{PP}} & \boldsymbol{\Sigma_{PD}} \\ \boldsymbol{\Sigma_{DP}} & \boldsymbol{\Sigma_{DD}} \end{pmatrix}$$

where  $\beta_{GP}$  is a vector of the SNP-telomere-length associations,  $\beta_{GD}$  is a vector of the SNP-disease associations,  $\beta_{IV}$  is the causal effect parameter, K is the number of SNPs,  $\Sigma_{PP}$  is a variance-covariance matrix with elements  $(\Sigma_{PP})_{ij} = se(\beta_{GPi})se(\beta_{GPj})\rho_{ij}$  where  $se(\beta_{GPi})$  is the standard error of the SNP-telomere-length association for the *i*th genetic variant, and  $\rho_{ij}$  is the correlation between the *i*th and *j*th variants due to linkage disequilibrium. Components of  $\Sigma_{DD}$  are similarly defined as  $(\Sigma_{DD})_{ij} = se(\beta_{GDi})se(\beta_{GDj})\rho_{ij}$ , and  $\Sigma_{PD} = \Sigma_{DP} = 0$  due to the two-sample setting (sensitivity analyses in a previous study<sup>13</sup> suggested results were robust to some correlation between the gene-phenotype and gene-outcome associations that may arise due to sample overlap). The slope estimated by maximum likelihood can be interpreted as the log odds ratio for disease per standard deviation change in genetically increased telomere length. The slope can further be interpreted as the causal effect of telomere length on disease if Mendelian randomization assumptions hold. The assumptions are: the SNPs are associated with telomere length (IV1); the SNPs are independent of confounders (IV2); and the SNPs are independent of disease adjusted for telomere length and confounders (IV3). See Supplementary Figure S7 for further details of the Mendelian randomization assumptions and Supplementary Table S6 for a glossary of terms.

199 The weighted median approach<sup>14</sup>

- 200 Let  $\hat{\beta}_{(1)},...,\hat{\beta}_{(J)}$  represent the J causal effect estimates ordered from smallest  $(\hat{\beta}_{(1)})$  to largest  $(\hat{\beta}_{(J)})$ .
- 201 Now define

202 
$$w_{(j)}^* = \frac{w_j}{S_J}$$
, where  $S_J = \sum_j w_j$ ,

- where  $w_j$  is the inverse variance of  $\hat{\beta}_{(j)}$ ,
- and equate  $\hat{\beta}_{(i)}$  with a quantile,  $p_{(i)}^{w}$ , defined as

205 
$$p_{(j)}^{w} = \frac{100}{S_{J}} \left( S_{(j)} - \frac{w_{(j)}}{2} \right).$$

- 206  $p_{(j)}^{w}$  represents the quantile from the weighted empirical distribution function of the ordered
- estimates  $\hat{\beta}_{(1)},...,\hat{\beta}_{(J)}$ . The weighted median estimate,  $\hat{\beta}_{WM}$  is defined as the  $50^{th}$  percentile of this
- weighted distribution. Typically the  $50^{th}$  percentile will lie between two estimates ( $\hat{\beta}_{(l)}$  and  $\hat{\beta}_{(m)}$ ,
- say), in which case  $\hat{\beta}_{WM}$  is found by linear interpolation.  $\hat{\beta}_{WM}$  is a consistent estimate for  $\beta$  provided
- 210 that at least 50% of the 'weight' making up  $S_J$  comes from genetic variants that are valid
- 211 instruments. In other words, the weighted median function provides a valid estimate of the causal
- 212 effect of telomere length on disease if at least half of the genetic information comes from valid
- 213 instruments (assumptions illustrated in Supplementary Figure S7; glossary of terms in
- 214 Supplementary Table S6). 14

- 216 The MR-Egger approach
- 217 The MR-Egger method<sup>15</sup> performs a weighted linear regression of the SNP-disease coefficients on
- the SNP-exposure coefficients (where exposure in this study is telomere length):

$$\frac{\hat{\Gamma}_{j}}{\sigma_{y_{j}}} = \frac{\beta_{0E}}{\sigma_{y_{j}}} + \beta_{1E} \frac{\hat{\gamma}_{j}}{\sigma_{y_{j}}}$$

where  $\Gamma$  corresponds to the SNP-disease coefficients,  $\gamma$  corresponds to the SNP-exposure coefficients and  $\sigma_{yj}$  is the standard error of  $\hat{\Gamma}_j$ . If all SNPs are valid instruments, then  $\beta_{0E}=0$ . The value of  $\hat{\beta}_{0E}$  can be interpreted as an estimate of the average pleiotropic effect across the SNPs. An intercept term that differs from zero is indicative of overall directional pleiotropy. The MR-Egger estimate for  $\beta$ ,  $\hat{\beta}_{1E}$ , is consistent even if *all* SNPs are invalid, provided that

- Across all SNPs, the magnitude of the SNP-exposure associations are independent of their pleiotropic effects (also known as the InSIDE [Instrument Strength Independent of Direct Effect] assumption)
- The number of SNPs, J, grows large (i.e. tends to infinity).

See Supplementary Figure S7 for further details on the assumptions and Supplementary Table S6 for a glossary of terms.

#### SUPPLEMENTARY RESULTS

In analyses of secondary cancer outcomes, genetically increased telomere length was associated with thyroid cancer, chronic lymphocytic leukemia and multiple myeloma (P<0.05) (Supplementary Figure S2). In analyses of secondary non-neoplastic diseases, genetically increased telomere length was associated with reduced odds of panic disorder (P<0.05) (Supplementary Figure S2). In secondary analyses of 44 risk factors for non-communicable diseases (Supplementary Table S2), genetically increased telomere length was associated with increased pulse pressure, systolic blood pressure, diastolic blood pressure, mean arterial pressure, triglycerides, uric acid and education and with decreased HDL cholesterol, mean corpuscular haemoglobin and mean corpuscular volume (P<0.05) (Supplementary Figure S5). There was some evidence for an association between genetically increased telomere length and ever smoking status (P=0.03, Supplementary Figure S6) but this association is unlikely to be reliable given that the SNP-telomere-length associations were adjusted for smoking history; the association may therefore reflect collider bias.<sup>21</sup>

#### SUPPLEMENTARY DISCUSSION

#### Mechanisms of association between SNPs and telomere length

The mechanisms of the underlying associations between the selected SNPs and telomere length are generally unknown. Some of the SNPs are located in or near the *TERC* or *TERT* genes, suggesting that the mechanism could involve the telomerase enzyme, as well as the *OBFC1* and *CTC1* genes, which have known roles in regulation of telomere length biology (Table 1). OBFC1 is an enzyme involved in initiating DNA replication and is involved in the telomere-associated CST complex. <sup>22</sup> *CTC1* encodes a component of the CST complex, which plays a role in protecting telomeres from degradation.

#### Bias from sample overlap and strength of the association between SNPs and telomere length

The selected genetic instruments for telomere length correspond to 10 independent genomic loci and collectively account for 2-3% of the variance in leukocyte telomere length. The corresponding F statistic is around 18, which means that bias due to weak instruments is unlikely to be substantial even if there were considerable overlap amongst the telomere length and disease and risk factor GWASs.<sup>23</sup> The estimated overlap in participants amongst the telomere length and outcome GWASs was less than 11% for all diseases and risk factors, except for hepatic steatosis, for which overlap was around 51%, indicating that the vast majority of our results should be robust to weak instrument bias.

#### Misconceptions about Mendelian randomization

A common misconception about Mendelian randomization studies is that genetic instruments should explain a substantial proportion of the variation in target exposures (e.g. telomere length in this study) in order to provide robust inferences about exposure-disease associations. However, if the genetic instruments are valid (i.e. conform to Mendelian randomization assumptions,

Supplementary Figure S7), the variation explained by the instrument only affects statistical power and does not generally affect validity of the causal inference. In this sense, genotype assignment in a Mendelian randomization study is analogous to treatment assignment in a randomized controlled trial, e.g. of blood pressure lowering drugs.<sup>24</sup> Although experimental interventions to reduce blood pressure may only explain a small fraction of the total variation in blood pressure in a typical RCT, we can still make causal inferences about blood pressure as a whole (and not just the proportion of variation in blood pressure due to the experimental intervention). Moreover, the aim of Mendelian randomization studies is to make inferences at the population level and not the individual level (for which genetic proxies of substantial explanatory power would be required).<sup>24</sup> If Mendelian randomization assumptions were violated, however, then the limited variation explained by our genetic instruments might not behave in similar manner to other sources of variation in telomere length, which would undermine our ability to draw causal inferences. See the above section 'Estimating the association between genetically increased telomere length and outcome traits' and Supplementary Figure S7 for details on the assumptions. See Supplementary Table S6 for an explanation of Mendelian randomization terminology. See Haycock et al<sup>25</sup> and Davey Smith and Hemani<sup>26</sup> for reviews on Mendelian randomization.

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#### Potential for confounding by population stratification, ancestry and age

It is unlikely that confounding by population stratification, ancestry or age (an important confounder of observational studies of telomere length) can account for our results. The 15 primary diseases showing some evidence of association with telomere length (defined as a P value<0.05) were 100% European, on the basis of self reported ancestry or genetic analyses (individuals showing genetic evidence of non-European ancestry were excluded). An addition, these studies all made some allowance for population stratification in their analyses: 12 adjusted for principal component scores of genetic variation in their models or applied genomic control corrections to their results; and 3 concluded there was little evidence for population stratification, on the basis of

visual inspection of Quantile-Quantile plots of GWAS results (i.e. lambdas for genomic inflation were close to 1). The GWAS we used to defined genetic instruments for telomere length<sup>4</sup> also adjusted for principal component scores; and lambdas for genomic inflation were close to 1. Since our MR analyses will have inherited any adjustments made in the original analyses, it is therefore unlikely that confounding by ancestry or population stratification can explain our results.

Confounding by age is also unlikely, given the random distribution of genotypes in the general population with respect to lifestyle and other environmental factors, as well as the fixed nature of germline genotypes. Consistent with this expectation, we did not observe an association between subject age and their genetically predicted telomere length values in our previous studies.<sup>44,45</sup>

#### Associations with non-neoplastic diseases

The inverse associations observed for coronary heart disease, abdominal aortic aneurysm, celiac disease and interstitial lung disease are compatible with findings based on observational and Mendelian randomization studies of telomere length as well as dyskeratosis congenita (a congenital disease characterized by chronically short telomeres). 46–50

**Supplementary Table S1.** Study characteristics for secondary non-communicable diseases and diseases from independent studies for replication analyses

	No.	No.	No. SNPs	Statistical	Pop.	First author /database
Cancer	cases	Controls	5111 5	power	1 ор.	1 list author/database
Chronic lymphocytic leukemia	2883	8350	1	0.22	EUR	Speedy/GWAS cat. <sup>51</sup>
Chronic myeloid leukemia	201	497	8	0.07	EA	Kim <sup>52</sup>
Ewing's sarcoma	401	684	4	0.06	EUR	Postel-Vinay <sup>53</sup>
Follicular lymphoma	212	748	3	0.04	EUR	Conde <sup>54</sup>
Gallbladder cancer	41	866	2	0.01	EA	Cha <sup>55</sup>
Gastric cancer						
Cardia adenocarcinoma	1126	2111	11	0.47	EA	Abnet <sup>56</sup>
Noncardia adenocarcinoma	632	2111	11	0.29	EA	Abnet <sup>56</sup>
Multiple myeloma	4692	10990	1	0.37	EUR	Chubb/GWAS cat.57
Nasopharyngeal carcinoma	1583	1894	2	0.17	EA	Bei <sup>58</sup>
B-cell Non-Hodgkin lymphoma	253	1438	10	0.13	EA	Tan <sup>59</sup>
Skin squamous cell carcinoma	449	11518	13	0.34	EUR	Zhang <sup>60</sup>
Thyroid cancer	649	431	12	0.16	EUR	Kohler <sup>61</sup>
Upper gastrointestinal cancers	3523	2100	2	0.28	EA	Li/dbGAP <sup>62</sup>
Autoimmune/inflammatory diseas	es					
Inflammatory psoriatic arthritis	609	990	13	0.29	EUR	Huffmeier <sup>63</sup>
Kawasaki disease	405	6252	11	0.26	EUR	Khor <sup>64</sup>
Narcolepsy	1188	1985	9	0.46	EA	Han <sup>65</sup>
Psoriasis	1139	1132	9	0.34	EA	Zhang <sup>66</sup>
Sarcoidosis	564	1575	9	0.16	EUR	Fischer <sup>67</sup>
Systemic lupus erythematosus	1311	1783	4	0.20	EUR	Hom/dbGAP <sup>68</sup>
Vitiligo	1117	1429	2	0.12	EA	Quan <sup>69</sup>
Wegener's granulomatosis	459	1503	10	0.20	EUR	Xie <sup>70</sup>
Neurological / psychiatric diseases						
Bulimia nervosa	151	2291	8	0.07	EUR	Wade <sup>71</sup>
Panic disorder	718	1717	8	0.28	EA	JCTGPD <sup>72</sup>
Parkinson's disease	1713	3978	4	0.35	EUR	Simón-Sánchez/dbGAP <sup>73</sup>

Other						
Hirschsprung's disease	173	615	6	0.04	EA	Tang <sup>74</sup>
Paget's disease	741	2699	12	0.43	EUR	Albagha <sup>75</sup>
Vascular dementia	84	200	8	0.03	EA	Kim <sup>76</sup>
Independent disease studies for						
Bladder cancer	7712	13125	1	0.56	EUR	Figueroa/GWAS cat. <sup>77</sup>
Colorectal cancer	728	3282	9	0.39	EA	Zhang <sup>78</sup>
Coronary heart disease	15399	15050	4	1.00	Mix	$C4D^{79}$
Glioma	1854	4955	1	0.12	EUR	GliomaScan/GWAS cat.80
Interstitial lung disease†	542	542	11	0.15	EUR	Noth <sup>81</sup>
Interstitial lung disease‡	242	1469	1	0.02	EA	Mushiroda/GWAS cat.82
Pancreatic cancer	4164	3792	10	0.90	EUR	PanC4 <sup>83</sup>
Multiple sclerosis	978	883	4	0.11	EUR	Baranzini/dbGAP <sup>84</sup>
Nasopharyngeal carcinoma	277	285	2	0.03	EA	Tse <sup>85</sup>
Type 2 diabetes	8569	8923	10	1.00	EA	Li <sup>86</sup>

Type 2 diabetes 8569 8923 10 1.00 EA Li<sup>∞</sup>

†≤17% cases overlapped with cases from Fingerlin et al<sup>51</sup> and 77% of cases had idiopathic pulmonary fibrosis; ‡all cases had idiopathic pulmonary fibrosis.

Study/database acronyms: C4D, Coronary Artery Disease Genetics Consortium; dbGAP, summary data downloaded from the database of Genotypes and Phenotypes; GWAS cat., data downloaded from the National Human Genome Research Institute/European Bioinformatics Institute Catalog of published genome wide association studies; JCTGPD, Japanese Collaboration Team for GWAS of Panic Disorder. Abbreviations: EUR, European; EA, East Asian; No., number; Pop., population; SNP, single nucleotide polymorphism.

Supplementary Table S2. Study characteristics of 44 risk factors for non-communicable diseases

							First
	Sample			No. of	Stat.		author /
	size	SD	Units	SNPs	power	Pop.	study
Anthropometric							
Birth length	22557	2.0	cm	12	1.00	EUR	$EGG^{87}$
Birth weight	26836	547.5	g	12	1.00	EUR	EGG <sup>88</sup>
Body mass index	241253	4.8	kg/m <sup>2</sup>	13	1.00	EUR	GIANT <sup>89</sup>
Childhood obesity	13848	NA	log <sub>e</sub> odds	12	0.78	EUR	$EGG^{90}$
Head circumference	10705	1.5	cm	13	1.00	EUR	$EGG^{91}$
Height	253288	0.1	m	13	1.00	EUR	GIANT <sup>92</sup>
Hip circumference	224459	8.5	cm	13	1.00	EUR	GIANT <sup>93</sup>
Waist circumference	224459	12.5	cm	13	1.00	EUR	GIANT <sup>93</sup>
Waist-to-hip ratio	224459	0.1	ratio	13	1.00	EUR	GIANT <sup>93</sup>
Smoking behaviors							
Age of smoking initiation	47961	0.3	log <sub>e</sub> years	13	1.00	EUR	TAG <sup>94</sup>
Cigarettes smoked per day	68028	11.7	CPD	13	1.00	EUR	$TAG^{94}$
Ever smoker	74035	NA	log <sub>e</sub> odds	13	1.00	EUR	$TAG^{94}$
Ex smoker	41969	NA	log <sub>e</sub> odds	13	1.00	EUR	TAG <sup>94</sup>
Blood pressure							
Diastolic blood pressure	66466	10.7	mm Hg	12	1.00	EUR	$ICBP^{95}$
Mean arterial pressure	27803	12.8	mm Hg	13	1.00	EUR	$ICBP^{96}$
Pulse pressure	70903	13.5	mm Hg	13	1.00	EUR	$ICBP^{96}$
Systolic blood pressure	66473	18.2	mm Hg	12	1.00	EUR	$ICBP^{95}$
Education			_				
College completion	95427	NA	log <sub>e</sub> odds	13	1.00	EUR	SSGAC <sup>97</sup>
Years of educational attainment	126559	1.2	years	13	1.00	EUR	SSGAC <sup>97</sup>
Glycemic							
2 hr glucose	15234	1.27	mmol/L	11	1.00	EUR	MAGIC <sup>98</sup>
Beta-cell function (HOMA-B)	46186	0.96	log <sub>e</sub> HOMA	12	1.00	EUR	MAGIC <sup>99</sup>
Fasting glucose	46186	0.73	mmol/L	12	1.00	EUR	MAGIC <sup>99</sup>
Fasting insulin	38238	0.79	log <sub>e</sub> pmol/L	12	1.00	EUR	MAGIC <sup>99</sup>

Fasting proinsulin	10701	0.81	log <sub>e</sub> pmol/L	12	1.00	EUR	MAGIC <sup>99</sup>
Gycated hemoglobin (HbA1c)	46368	0.53	%	12	1.00	EUR	MAGIC <sup>100</sup>
Insulin resistance (HOMA-IR)	46186	0.67	$\log_{\mathrm{e}}\mathrm{HOMA}$	12	1.00	EUR	MAGIC <sup>99</sup>
Hemotological							van der
Hemoglobin	54287	1.3	g/dL	12	1.00	EUR	Harst <sup>101</sup>
Mean cell hemoglobin	45969	1.99	pg	12	1.00	EUR	van der Harst <sup>101</sup>
Mean cell hemoglobin concentration	49632	1.01	g/dL	12	1.00	EUR	van der Harst <sup>101</sup>
Mean cell volume	51277	5.2	fl	12	1.00	EUR	van der Harst <sup>101</sup>
Packed cell volume	46848	5.9	%	12	1.00	EUR	van der Harst <sup>101</sup>
Red blood cell count	47873	0.5	$10^{12}/L$	12	1.00	EUR	van der Harst <sup>101</sup>
Lipids							102
HDL cholesterol	103019	15.51	mg/dL	11	1.00	EUR	GLGC <sup>102</sup>
LDL cholesterol	97562	38.67	mg/dL	11	1.00	EUR	$GLGC^{102}$ $GLGC^{102}$
Total cholesterol	103266 99050	41.75 90.72	mg/dL mg/dL	11 11	1.00 1.00	EUR EUR	GLGC <sup>102</sup>
Triglycerides Renal function	99030	90.72	mg/uL	11	1.00	LUK	GLGC
Kenai function							CKDGen <sup>10</sup>
Microalbuminuria	30482	NA	log <sub>e</sub> odds	13	0.82	EUR	CKDGen <sup>10</sup>
Serum creatinine	67093	0.24	$log_eml/min/1.73m^2$	13	1.00	EUR	CKDGen <sup>10</sup>
Serum cystatin	20957	0.23	log <sub>e</sub> ml/min/1.73m <sup>2</sup>	13	1.00	EUR	3
Urinary albumin-to-creatinine ratio	31580	1.0	log <sub>e</sub> mg/g	13	1.00	EUR	CKDGen <sup>10</sup>
Other Grade of nuclear cataract	7140	0.8	grade	11	1.00	ASN	SEEDS <sup>104</sup>
Hepatic steatosis	7176	5.6	Hounsfield units	12	1.00	EUR	Speliotes <sup>10</sup>
Percent emphysema	7914	0.71	$\log_e \% + 1$	12	1.00	ME	MESA <sup>106</sup>
Uric acid	42742	1.3	mg/dL	12	1.00	EUR	GUGC <sup>107</sup>

Study acronyms: CKDGen, chronic kidney disease genetics consortium; EGG, Early Growth Genetics Consortium; GIANT, Genetic Investigation of ANthropometric Traits; GUGC, Global Urate and Gout consortium; TAG, Tobacco and Genetics Consortium; ICBP, International Consortium for Blood Pressure; SSGAC, Social Science Genetics Association Consortium; MAGIC, Meta-Analyses of Glucose and Insulin-related traits Consortium; MESA, Multi-Ethnic Study of Atherosclerosis; GLGC, Global Lipids Genetics Consortium; SEEDS, the Singapore Epidemiology of Eye Diseases Study.

Abbreviations: ASN, Asian; Con., concentration; EUR, European population; ME, multi-ethnic; SD - standard deviation; loge, natural log; Stat., statistical

Supplementary Table S3. Selected prospective observational studies of the association between leukocyte telomere length and disease

				NT C					DD (050/				
				No. of controls	No.	RR (95% CI)	Scale of RR		RR (95% CI) per SD				
Cohort / first				/ cohort	of	as reported by	reported by	Conversion	increase in				Search
author	Disease	Year	Design	size	cases	study	study	factor§	TL	Adjusted <sup>‡</sup>	Pop.	$P_{het}$	strategy†
Cancer outco	mes					·	<u>,                                      </u>			,			25
NHS, HPFS <sup>108</sup>	Bladder cancer	2007	NCC	192	184	1.88 (1.05 to 3.36)	shortest vs. longest quartile	2.54	1.28 (1.02 to 1.61)	++	EUR	NA	2
CCHS, CGPS <sup>109</sup>	Breast cancer	2013	PC	24588	574	0.99 (0.95 to 1.03)	per 1000 bp (1.29 SD) decrease	-1.29	1.01 (0.98 to 1.04)	+++++	EUR		1
SWHS <sup>110</sup>	Breast cancer	2013	NCC	695	601	1.77 (1.02 to 3.06)	shortest vs. longest quintile	2.80	1.23 (1.01 to 1.49)	++	EA	0.17	2
Sister Study <sup>111</sup>	Breast cancer	2011	Case- cohort	735	342	0.93 (0.64 to 1.35)	shortest vs. longest quartile	-2.54	1.03 (0.89 to 1.19)	+	EUR (92%)	0.17	1
EPIC <sup>112</sup>	Breast cancer	2010	NCC	420	199	1.58 (0.75 to 3.31)	shortest vs. longest quartile	2.54	1.2 (0.89 to 1.6)	+	EUR		1
WHS <sup>113</sup>	Colorectal cancer	2010	NCC	357	134	0.94 (0.65 to 1.38)	per unit (1.30 SD) decrease	-1.30	1.05 (0.78 to 1.4)	+++++	EUR		3
PHS <sup>114</sup>	Colorectal cancer	2009	NCC	306	191	0.8 (0.55 to 1.16)	per unit (1.72 SD) decrease	-1.72	1.14 (0.92 to 1.41)	++++	EUR		3
CCHS, CGPS <sup>109</sup>	Colorectal cancer	2013	PC	46748	496	0.97 (0.88 to 1.07)	per 1000 bp (1.29 SD) decrease	-1.29	1.02 (0.95 to 1.1)	++++	EUR	0.47	1
SWHS <sup>115</sup>	Colorectal cancer	2012	NCC	549	441	1.61 (0.94 to 2.75)	longest vs.  3rd shortest quintile	1.40	1.4 (0.96 to 2.06)	+	EA		1
EPIC <sup>112</sup>	Colorectal cancer	2010	NCC	406	185	1.13 (0.54 to 2.36)	shortest vs. longest quartile	-2.54	0.95 (0.71 to 1.27)	+	EUR		1
NHS <sup>116</sup>	Endometrial cancer	2010	NCC	791	279	1.2 (0.73 to 1.96)	shortest vs. longest quartile	-2.54	0.93 (0.77 to 1.13)	+++++	EUR	0.11	2
CCHS, CGPS <sup>109</sup>	Endometrial cancer	2013	PC	25262	103	0.85 (0.71 to 1.02)	per 1000 bp (1.29 SD)	-1.29	1.13 (0.99 to 1.31)	+++++	EUR	*	1

decrease

PLCO <sup>117</sup>	Glioma	2013	NCC	198	101	1.26 (0.69 to 2.29)	shortest vs. longest tertile	-2.18	0.9 (0.68 to 1.18)	++	EUR	NA	1
CCHS, CGPS <sup>109</sup>	Head & neck cancer	2013	PC	47036	76	1.17 (0.9 to 1.53)	per 1000 bp (1.29 SD) decrease	-1.29	0.89 (0.72 to 1.09)	++++	EUR	NA	1
CCHS, CGPS <sup>109</sup>	Kidney cancer	2013	PC	47063	59	1.04 (0.78 to 1.39)	per 1000 bp (1.29 SD) decrease	-1.29	0.97 (0.77 to 1.21)	++++	EUR	NA	1
PLCO <sup>118</sup>	Kidney cancer	2013	NCC	410	209	0.8 (0.5 to 1.5)	longest vs. shortest quartile	2.54	0.92 (0.74 to 1.14)	+++	EUR (89.5%)	NA	1
PLCO, ATBC, SWHS <sup>119</sup>	Lung adenocarcinoma	2014	NCC	288	288	2.52 (1.38 to 4.6)	longest vs. shortest quartile	2.54	1.44 (1.14 to 1.82)	++	EUR (75%)	NA	1
CCHS, CGPS <sup>109</sup>	Lung cancer	2013	PC	47035	522	1.08 (0.98 to 1.2)	per 1000 bp (1.29 SD) decrease	-1.29	0.94 (0.87 to 1.02)	++++	EUR	< 0.001	1
PLCO, ATBC, SWHS <sup>119</sup>	Lung cancer	2014	NCC	847	847	1.86 (1.33 to 2.62)	longest vs. shortest quartile	2.54	1.28 (1.12 to 1.46)	++	EUR (75%)	<0.001	1
PLCO, ATBC, SWHS <sup>119</sup>	Lung SCC	2014	NCC	163	163	1.14 (0.53 to 2.45)	longest vs. shortest quartile	2.54	1.05 (0.78 to 1.42)	++	EUR (75%)	NA	1
CCHS, CGPS <sup>109</sup>	Melanoma	2013	PC	46805	177	0.89 (0.77 to 1.03)	per 1000 bp (1.29 SD) decrease	-1.29	1.09 (0.98 to 1.23)	++++	EUR	0.02	1
WHI, HPFS, NHS <sup>120</sup>	Melanoma	2011	NCC	579	557	0.43 (0.27 to 0.7)	shortest vs. longest quartile	-2.54	1.39 (1.16 to 1.68)	+	EUR	0.03	2
CCHS, CGPS <sup>109</sup>	Ovarian cancer	2013	PC	25367	96	0.85 (0.7 to 1.03)	per 1000 bp (1.29 SD) decrease	-1.29	1.13 (0.98 to 1.32)	+++++	EUR	NA	1
CCHS, CGPS <sup>109</sup>	Pancreatic cancer	2013	PC	47091	124	1.14 (0.93 to 1.41)	per 1000 bp (1.29 SD) decrease	-1.29	0.9 (0.77 to 1.06)	++++	EUR		1
ATBC <sup>121</sup>	Pancreatic cancer	2013	NCC	660	193	1.58 (1.02 to 2.46)	longest vs. shortest quartile	2.54	1.2 (1.01 to 1.42)	++	EUR	0.05	1

EPIC <sup>122</sup>	Pancreatic cancer	2014	NCC	331	331	1.38 (0.8 to 2.41)	longest vs. shortest quartile	2.54	1.13 (0.91 to 1.41)	+	EUR		1
CCHS, CGPS <sup>109</sup>	Prostate cancer	2013	PC	21387	418	0.94 (0.85 to 1.04)	per 1000 bp (1.29 SD) decrease	-1.29	1.05 (0.97 to 1.13)	++++	EUR	0.37	1
HPFS <sup>123</sup>	Prostate cancer	2015	NCC	935	922	1.11 (1.01 to 1.22)	per SD increase	1.00	1.11 (1.01 to 1.22)	++++	EUR		1
NHS <sup>124</sup>	Skin BCC	2011	NCC	1683	363	0.91 (0.66 to 1.25)	longest vs. shortest quartile	2.54	0.96 (0.85 to 1.09)	+	EUR	NA	1
CCHS, CGPS <sup>109</sup>	Testicular cancer	2013	PC	21568	10	1.09 (0.57 to 2.09)	per 1000 bp (1.29 SD) decrease	-1.29	0.94 (0.56 to 1.55)	++++	EUR	NA	1
Non-neoplast	ic diseases												
$Haycock^{\parallel 125}$	Coronary heart disease	2014	MA	27352	2272	1.4 (1.15 to 1.7)	shortest vs. longest tertile	-2.18	0.86 (0.78 to 0.94)	*	EUR	NA	4
Haycock <sup>#125</sup>	Ischemic stroke	2014	MA	5300	824	1.14 (0.85 to 1.54)	shortest vs. longest tertile	-2.18	0.94 (0.82 to 1.08)	*	EUR	NA	4
Bruneck, SHFS, WHI <sup>126</sup>	Type 2 diabetes	2014	MA	6991	2011	1.31 (1.07 to 1.6)	shortest vs. longest quartile	-2.54	0.9 (0.83 to 0.97)	**	Mix	NA	4

†Search strategy used to identify the study (see Table S4 for details). Meta-analysis of 11 prospective studies; "Meta-analysis of 6 prospective studies (90% of cases were ischemic stroke, 10% were unclassified cerebrovascular disease); \*To convert reported log RR to log RR per SD increase in telomere length; \*Adjustment for confounders: +adjusted for age and sex; ++plus smoking; +++plus body mass index; +++++plus alcohol and/or physical activity; +++++plus hormone replacement therapy, menopause and/or parity; \*most studies adjusted for age, sex and non-lipid vascular risk factors; \*\*adjusted for age, sex and body mass index.

Acronyms/abbreviations: BCC, basal cell carcinoma; bp, base pairs; CI, confidence interval; EA, East Asian; EUR, European; MA, random-effects meta-analysis of prospective studies; NCC, nested case-control study;
PC, prospective cohort; Phet, p value for heterogeneity between studies; Pop., population; RR, relative risk; SD, standard deviation; SCC, squamous cell carcinoma; vs., versus; TL, telomere length. Study acronyms:
ATBC, Alpha-Tocopherol, Beta-Carotene Cancer Prevention Study; CCHS, Copenhagen General Population Study; EPIC, European Prospective Investigation into Cancer and
Nutrition study; HPFS, Health Professionals Follow-Up Study; NHS, Nurses Health Study; PLCO, Prostate, Lung, Colorectal, and Ovarian; SHFS, Strong Heart Family Study; the Sister
Study: SWHS, Shanghai Women's Health Study; WHI, Women's Health Initiative; WHS, Women's Health Study

# Supplementary Table S4. PubMed search strategy for prospective observational studies of association between telomere length\* and disease

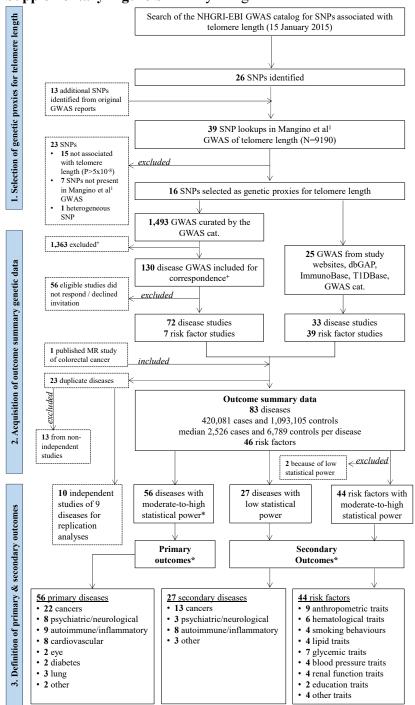
		No. of	No. meeting	Reasons	No. of
Search		studies	inclusion	for further	studies
strategy	Search terms or meta-analysis	identified	criteria	exclusions	included
Inclusion cri	teria: prospective study of primary cancer outcome and telomere length†				
Strategy 1	25 February 2015: cancer[TIAB] AND telomere length[TIAB] AND (meta analysis[TIAB] OR prospective[TIAB] OR meta-analysis[TIAB])	54	11	NA	11 <sup>‡</sup>
Strategy 2	25 March 2015: telomere length[Title/Abstract] AND (retrospective[Title/Abstract] OR case-control[Title/Abstract] OR case control[Title/Abstract] OR meta-analysis[Title/Abstract] OR prospective[Title/Abstract] OR cohort[Title/Abstract] OR cross-sectional[Title/Abstract] OR cross-sectional[Title/Abstract] OR cross-sectional[Title/Abstract] OR breast cancer[Title/Abstract] OR chronic myeloid leukemia[Title/Abstract] OR esophageal adenocarcinoma[Title/Abstract] OR endometrial cancer[Title/Abstract] OR esophageal cancer[Title/Abstract] OR gallbladder cancer[Title/Abstract] OR glioma[Title/Abstract] OR head cancer[Title/Abstract] OR neck cancer[Title/Abstract] OR oesophageal adenocarcinoma[Title/Abstract] OR kidney cancer[Title/Abstract] OR melanoma[Title/Abstract] OR nasopharyngeal carcinoma[Title/Abstract] OR neuroblastoma[Title/Abstract] OR non-melanoma skin cancer[Title/Abstract] OR basal cell carcinoma[Title/Abstract] OR squamous cell carcinoma[Title/Abstract] OR ovarian cancer[Title/Abstract] OR pancreatic cancer[Title/Abstract] OR prostate cancer[Title/Abstract] OR testicular germ cell cancer[Title/Abstract] OR Wilm's tumour[Title/Abstract] OR Bladder cancer[Title/Abstract] OR Breast cancer[Title/Abstract] OR Chronic lymphocytic leukemia[Title/Abstract] OR Colorectal cancer[Title/Abstract] OR Multiple myeloma[Title/Abstract] OR Lung squamous cell cancer[Title/Abstract] OR cancer[Title/Abstract] OR leukemia[Title/Abstract] OR leukemi	209	17	13 duplicates	4
Strategy 3	Ma et al <sup>127</sup> (2011) and Wentzensen et al <sup>128</sup> (2011)	48	10	8 duplicates	2
Inclusion cri	teria: prospective study of primary disease outcome and telomere length†			•	
Strategy 4	8 January 2016: (meta-analysis OR "meta analysis") AND "telomere length"	42	7	2 did not report relative risks <sup>§</sup> ; 3 duplicates	2 <sup>  </sup>

<sup>\*</sup>all identified eligible studies were studies of leukocyte telomere length; †1 study reported findings for 2 primary cancer outcomes and 1 study reported findings for 11 primary cancer outcomes; |1 meta-analysis reported findings for 2 primary non-neoplastic diseases; †primary outcomes were diseases where a priori statistical power was >50% to detect associations with telomere length (see supplementary text for technical details); see table S1 for a list of the primary disease outcomes; §relative risks were defined as odds ratios, hazard ratios and risk ratios

# Supplementary Table S6. Glossary of terms

Mendelian randomization	A technique to appraise causality in observational studies using
	genetic variants as 'unconfounded' instruments for risk factors or
	modifiable exposures of interest.
Instrumental variable	A 'proxy' variable used in place of the hypothesized risk factor
	or exposure in a Mendelian randomization analysis. A valid
	instrumental variable is associated with the exposure of interest
	but is not associated with confounders; and is associated with the
	outcome (e.g. disease) exclusively via its effect on the
	hypothesized exposure (see Supplementary Figure S7 for an
	illustration of these assumptions).
Reverse causation	When the outcome causes variation in the hypothesized exposure
	and not <i>vice versa</i> .
Confounding	When the association between exposure and outcome is not due
	to a causal relationship between the two variables but arises as a
	result of the separate effects of a third variable (the confounder)
	on the exposure and the outcome. Mendelian randomization
	studies are less susceptible to confounding in comparison to
	observational studies (but confounding by pleiotropy or
	population stratification is possible).
Pleiotropy	Occurs when a genetic variant is associated with multiple traits or
	phenotypes. Vertical pleiotropy occurs when the phenotypes are
	on the same causal pathway (and is less problematic for
	Mendelian randomization studies). Horizontal pleiotropy occurs
	if the phenotypes are associated with the genetic variant via
	separate pathways and can introduce confounding into a
	Mendelian randomization analysis. Sensitivity analyses, such as
	MR-Egger, the weighted median, scatter plots and funnel plots,
Collider bias	can be used to test and, in some instances, adjust for pleiotropy.
Conider bias	The phenomenon by which statistical adjustment for a variable,
	M (known as the collider), that is a downstream consequence of both the exposure X and the outcome Y, induces an association
	between X and Y that was not previously present, and therefore
	leads to bias. In MR, if published genetic associations with the
	exposure and/or outcome are adjusted for a collider, this may
	lead to collider bias.
Weak instrument bias	Occurs when the instrument is only weakly associated with the
Weak instrument olds	exposure. Can introduce confounding into a Mendelian
	randomization analysis when the exposure and outcome data
	come from the same sample. When exposure and outcome data
	come from separate samples, as in two-sample Mendelian
	randomization, bias is towards the null. An F statistic > 10, for
	the association between the instrument and exposure, is
	sometimes used as a threshold for defining strong instruments,
	although weak instrument bias varies continuously with the
	strength of the F statistic.

## 335 Supplementary Figure S1. Study design



+We searched the GWAS catalog in January 2015 for studies of non-communicable diseases that did not select controls on the basis of pre-existing conditions. Of the 1493 studies in the GWAS catalog with unique PubMed reference numbers, we classified 773 as disease studies (the excluded nondisease studies were typically studies of risk factors for disease, biomarkers or response to treatments). A further 103 studies were excluded for the following reasons: studies of infectious diseases, studies of congenital abnormalities, studies of (not-cause specific) mortality, studies nested within disease populations and studies using pooled DNA samples. Of the 670 remaining noncommunicable disease studies, 130 were identified for correspondence. Our objective was to obtain the single largest available study for each non-communicable disease, so as to avoid unnecessary correspondence with duplicate studies and to avoid including studies with overlapping samples. \*Primary outcomes were diseases with sufficient cases and controls for >50% power and secondary outcomes were diseases with <50% power to detect odds ratios ≥2.0 per standard deviation change in genetically increased telomere length (alpha assumed to be 0.01). All risk factors were classified as secondary outcomes. GWAS, genome-wide association study; GWAS Cat., NHGRI-EBI GWAS catalogue; SNP, single nucleotide polymorphism; NHGRI, National Human Genome Research Institute; EBI, European Bioinformatics Institute

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<sup>\*</sup>P value for association between genetically increased telomere length and disease from maximum likelihood; Pheb, P value for heterogeneity amongst SNPs within the genetic risk score; SNP, single nucleotide polymorphism; CI, confidence interval

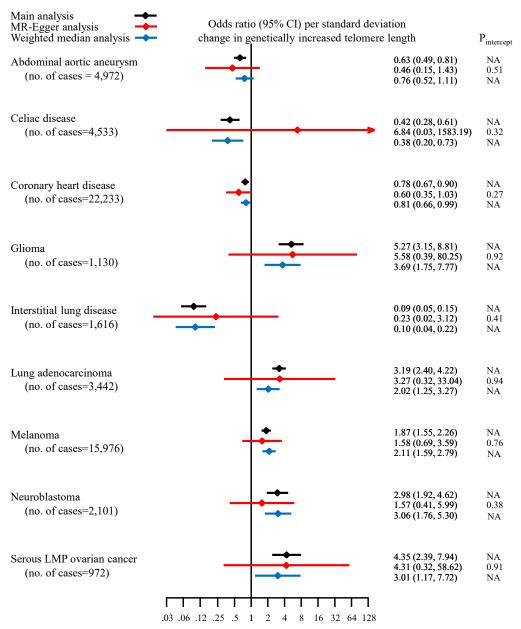
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# **Supplementary Figure S3**. Replication of association between genetically increased telomere length and odds of non-communicable diseases in independent datasets

Disease/ Study	No. of cases	No. of SNPs		ratio (95% CI) per standard deviation change in genetically increased telomere length		P*	$P_{het}$
Coronary heart dise	ase						
CARDIoGRAM†	22233	13	+		0.78 (0.67, 0.90)	0.0009	0.2441
C4D	15399	4	<b>-</b>		0.70 (0.56, 0.88)	.0023	0.0668
Colorectal cancer							
CORECT/GECCO†		9		<del>*</del>	1.09 (0.91, 1.31)	0.3436	0.016
Zhang et al	728	7	_	<del>  •                                     </del>	1.29 (0.64, 2.59)	.4738	0.2690
Multiple sclerosis							
IMSGC†	14498	3	_	<del>(_</del>	0.98 (0.70, 1.36)	0.8885	0.1455
Baranzini et al	978	4			1.09 (0.47, 2.55)	.8444	0.4628
Type 2 diabetes							
DIAGRAM†	10415	11	-	<b>†</b> .	1.00 (0.84, 1.20)	0.9837	0.6811
Li et al	8569	8	_		1.28 (0.77, 2.11)	.3407	0.8439
Bladder cancer							
NBCS†	1601	10		<b>—</b>	2.19 (1.32, 3.66)	0.0026	0.2541
Figueroa et al	7712	1			5.21 (2.48, 10.94)	1.00x10 <sup>-5</sup>	NA
Pancreatic cancer							
PanScan†	5105	12	<b>—</b>	╆	0.86 (0.56, 1.32)	0.5009	0.0016
PanC4	4164	11	-		0.74 (0.53, 1.02)	0.0657	0.0435
Glioma							
Walsh et al†	1130	12		<b>—</b>	5.27 (3.15, 8.81)	2.45x10 <sup>-10</sup>	
GliomaScan	1854	1		<del></del>	21.55 (3.82, 121.47)	0.0005	NA
Interstitial lung dise	ase						
Fingerlin et al†	1616	9 —			0.09 (0.05, 0.15)	2.02x10 <sup>-19</sup>	0.0014
Noth et al+	542	11	<b>—</b>		0.30 (0.12, 0.77)	0.0120	0.1833
Nasopharyngeal car							
Bei et al†	1583	2	_	<del>  •</del>	1.28 (0.59, 2.76)	0.5348	0.1200
Tse et al	277	2		<b>†</b>	5.04 (0.36, 71.44)	0.2315	0.1659
		.06 .12	.25 .5	1 2 4 8 16 32 64			

\*P value for association between genetically increased telomere length and disease from maximum likelihood. †Primary or secondary study from Fig. 1 or Fig. S2. \*Noth et al<sup>81</sup>: ≤17% of the cases overlapped with cases from Fingerlin et al<sup>31</sup> and 77% of cases had idiopathic pulmonary fibrosis; ‡An inverse association was also observed in Mushiroda et al<sup>82</sup>. P<sub>het</sub>, p value for heterogeneity amongst SNPs in the genetic risk score (NA when only a single SNP available); SNP, single nucleotide polymorphism; C1, confidence interval. **Study abbreviations:** C4D, Coronary Artery Disease Genetics Consortium; C4RDIoGRAM, Coronary Artery Disease Genome wide Replication and Meta-analysis; CORECT, ColoRectal Transdisciplinary Study; GECCO, Genetics and Epidemiology of Colorectal Cancer Consortium; IMSGC, International Multiple Sclerosis Genetic Consortium; NBCS, Nijmegen Bladder Cancer Study; IMSGC, International Multiple Sclerosis Genetic Consortium.

**Supplementary Figure S4**. Sensitivity analyses of association between genetically increased telomere length and odds of non-communicable diseases



LMP, low malignancy potential; CI, confidence interval. The  $P_{intercept}$  from MR-Egger regression tests the null hypothesis that the intercept is zero and can be interpreted as a statistical test for the presence of directional (bias inducing) pleiotropy; the smaller the  $P_{intercept}$  value the stronger the evidence for directional pleiotropy.

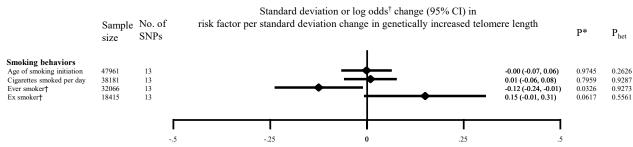
# Supplementary Figure S5. Association between genetically increased telomere length and risk factors for non-communicable diseases Sample No. of Standard deviation or log odds<sup>1</sup> change (95% CD) in risk factor

	Sample size	No. of SNPs		igth	P*	$P_{het}$
Anthropometric traits Height Body mass index Waist circumference Hip circumference Hip circumference Waist-to-hip ratio Birth weight Birth length Childhood obesity† Head circumference	247695 241253 158648 149224 148662 26836 22557 13848 10705	13 13 13 13 13 12 12 12 12		0.02 (-0.01, 0.05) -0.01 (-0.04, 0.03) 0.01 (-0.04, 0.05) -0.00 (-0.05, 0.04) 0.02 (-0.02, 0.06) 0.00 (-0.08, 0.08) -0.05 (-0.15, 0.04) 0.16 (-0.10, 0.43) -0.06 (-0.20, 0.09)	0.2477 0.6054 0.7911 0.8472 0.3158 0.9708 0.2753 0.2286 0.4416	<0.0001 0.1109 0.1302 0.1708 0.2823 0.6970 0.9138 0.2111 0.2177
Education Years of educational attainment College completion†	126559 126559	13 13	<del>-</del>	0.04 (0.01, 0.07) 0.12 (0.02, 0.21)	0.0142 0.0215	0.4718 0.1764
Lipids Total cholesterol HDL cholesterol Triglycerides LDL cholesterol	103266 103019 99050 97562	11 11 11 11		-0.00 (-0.05, 0.05) -0.08 (-0.13, -0.04) 0.07 (0.03, 0.12) 0.00 (-0.05, 0.05)	0.9899 0.0005 0.0012 0.9985	0.0037 0.2924 0.4907 0.0294
Blood pressure Pulse pressure Systolic blood pressure Diastolic blood pressure Mean arterial pressure	70903 66473 66466 27803	13 12 12 13	<del>+</del>	0.06 (0.01, 0.10) 0.09 (0.04, 0.15) 0.10 (0.04, 0.16) 0.09 (0.04, 0.13)	0.0148 0.0014 0.0008 0.0005	0.1526 0.2368 0.6963 0.2146
Renal function Serum creatinine Urinary albumin-to-creatinine ratio Microalbuminuria† Serum cystatin	67093 31580 30482 20957	13 13 13 13		0.02 (-0.03, 0.07) 0.09 (-0.00, 0.19) 0.20 (-0.06, 0.46) 0.02 (-0.07, 0.12)	0.4843 0.0546 0.1308 0.6247	0.2522 0.2306 0.5607 0.4767
Hemotological traits Hemoglobin Mean cell volume Mean cell hemoglobin concentration Red blood cell count Packed cell volume Mean cell hemoglobin	54287 51277 49632 47873 46848 45969	12 12 12 12 12 12	<del></del>	-0.01 (-0.05, 0.04) -0.09 (-0.14, -0.04) -0.01 (-0.03, 0.01) 0.03 (-0.01, 0.08) -0.00 (-0.03, 0.03) -0.23 (-0.34, -0.12)	0.7553 0.0009 0.3332 0.1626 0.8309 <0.0001	0.6636 0.0062 0.1728 0.4471 0.4526 0.0160
Glycemic traits Gycated hemoglobin (HbA1e) Fasted glucose Fasted insulin Insulin resistance (HOMA-IR) Beta-cell function (HOMA-B) 2hr glucose Fasted proinsulin	46368 46186 46186 46186 46186 15234 10701	12 12 12 12 12 12 11 12		-0.01 (-0.07, 0.05) 0.01 (-0.04, 0.06) -0.05 (-0.10, 0.00) -0.05 (-0.11, 0.01) -0.03 (-0.06, 0.01) -0.12 (-0.27, 0.02) 0.06 (-0.03, 0.15)	0.7766 0.6798 0.0586 0.1259 0.1779 0.1016 0.2139	0.3652 0.2955 0.1910 0.2511 0.0165 0.9574 0.8945
Other traits Uric acid Percent emphysema Hepatic steatosis Grade of nuclear cataract	42742 7914 7176 7140	12 12 12 12 8	-5 -25 <b>0</b> .25 .5	0.02 (0.00, 0.03) 0.09 (-0.04, 0.23) 0.11 (-0.08, 0.29) -0.00 (-0.15, 0.14)	0.0341 0.1826 0.2651 0.9572	0.0015 0.5247 0.8700 0.1934

<sup>\*</sup>P value for association between genetically increased telomere length and risk factor from maximum likelihood;  $P_{het}$ , p value for heterogeneity amongst SNPs within the genetic risk score; SNP, single nucleotide polymorphism; CI, confidence interval; HbA1c, hemoglobin A1c; HOMA-B, homeostatic model assessment  $\beta$ -cell function; IR, insulin resistance;  $\dagger$  for binary risk factors results reflect the log odds ratio for the risk factor, all other results reflect the standard deviation change in the risk factor

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# Supplementary Figure S6. Association between genetically increased telomere length and smoking

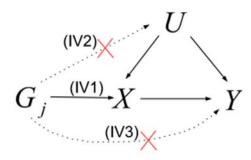


<sup>\*</sup>P value for association between genetically increased telomere length and risk factor from maximum likelihood; Phet, P value for heterogeneity amongst SNPs within the genetic risk score; SNP, single nucleotide polymorphism; CI, confidence interval; †for binary risk factors results reflect the log odds ratio for the risk factor, all other results reflect the standard deviation change in the risk factor

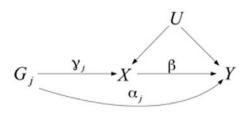
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**Supplementary Figure S7.** Causal diagram illustrating the assumptions of Mendelian randomization

368 a)



370 b)



IV, instrumental variable assumption;  $G_j$ , single nucleotide polymorphism j; X, telomere length; Y, outcome (disease or risk factor); U, confounder;  $\alpha$ , G-Y association not mediated by telomere length (often described as a horizontal pleiotropic or direct effect);  $\gamma$ , SNP-telomere-length association.

a) Key assumptions of Mendelian randomization.  $G_j$  is associated with X (IV1);  $G_i$  is independent of confounders (IV2);  $G_i$  is independent of Y given X and U (IV3). The weighted median approach assumes that IV1-IV3 hold for genetic variants making up at least 50% of the weight in the analysis; MR-Egger relaxes assumption IV3 (see InSIDE assumption below).

**b)** Assumptions underlying the MR-Egger approach. IV3 is replaced with the InSIDE assumption (Instrument Strength Independent of Direct Effect): the strength of the pleiotropic effect ( $\alpha$ j) does not correlate with the strength of the G-X association ( $\gamma$ j). Under the InSIDE assumption, MR-Egger can consistently estimate the causal effect of X on Y, represented by the parameter  $\beta$  in (b).

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with 397 cases and 384 controls): The Vascular Research Consortium of New Zealand recruited

New Zealand men and women with a proven history of AAA (infra-renal aortic diameter ≥ 30 mm proven on ultrasound or CT scan). Approximately 80% had undergone surgical AAA repair (typically AAA's > 50-55 mm in diameter). The vast majority of cases (>97%) were of Anglo-European ancestry. The control group underwent an abdominal ultrasound scan to exclude (>25 mm) concurrent abdominal aortic aneurysm and Anglo-European ancestry was required for inclusion. Controls were also screened for peripheral artery disease (PAD; using ankle brachial index), carotid artery disease (ultrasound) and other cardiovascular risk factors.

Geisinger Vascular Clinic AAA Study, Pennsylvania, USA: AAA patients (n=724) were enrolled through the Department of Vascular Surgery at Geisinger Medical Center, Danville, PA. Details of this case-control set have been reported previously, and the samples have been used in previous association studies. To identify cases and controls from the electronic medical records, an ePhenotyping algorithm was developed. AAA cases were defined as infrarenal aortic diameter ≥ 30 mm as revealed by abdominal imaging. Approximately 20% of individuals with AAA had a family history of AAA. A control group (n=1231) was obtained through the Geisinger MyCode. Project, a cohort of Geisinger Clinic patients recruited for genomic studies. The MyCode. controls were matched for age distribution and sex to the Geisinger Vascular Clinic AAA cases. Based on electronic medical records, controls had no ICD-9 codes for AAA in their records, but they were not screened by ultrasonography for AAA. Both cases and controls from the Geisinger Clinic were of European descent. The eMERGE Network Imputed GWAS for 41 Phenotypes (the dbGaP eMERGE Phase 1 and 2 Merged data Submission) accession number is: phs000888.v1.p1 which includes the Geisinger AAA data.

Iceland, deCODE Genetics: Icelandic individuals with AAA (defined as infra-renal aortic diameter ≥ 30 mm) were recruited from a registry of individuals who were admitted at Landspitali University Hospital, in Reykjavik, Iceland, 1980 – 2006. AAA patients were either followed up or treated by

intervention for emergency repair of symptomatic or ruptured AAA or for an elective repair by surgery or endovascular intervention. In total, whole genome data from 557 subjects with AAA, enrolled as part of the CVD genetics program at deCODE, were included in the metaGWAS. The Icelandic controls used (n=89,235) were selected from among individuals who have participated in various GWA studies and who were recruited as part of genetic programs at deCODE. Individuals with known cardiovascular disease were excluded as controls<sup>129</sup> but controls were unscreened for AAA.

The Netherlands: The AAA sample set from Utrecht was recruited in 2007-2009 from eight centres in The Netherland<sup>129</sup>, mainly when individuals visited their vascular surgeon in the polyclinic or, in rare cases, during hospital admission for elective or emergency AAA surgery. An AAA was defined as an infrarenal aorta ≥ 30 mm. The sample set (n=840) comprised 89.9% males, with a mean AAA diameter of 58.4 mm, 61.7% had received surgery, of which 8.1 % was after rupture. The Dutch controls (n=2791) used in the AAA GWAS were recruited as part of the Nijmegen Biomedical Study and the Nijmegen Bladder Cancer Study (see http://dceg.cancer.gov/icbc/membership.html).

Meta-analysis of AAA GWASs

Data from the six cohorts detailed above, comprising 4972 AAA cases and 99,858 controls, that were genotyped with a variety of genome-wide SNP arrays. All cohorts underwent quality control filtering using the manufacturers' array-specific guidelines but with consistently applied inclusion criteria of SNP or sample call rates >95% and Hardy-Weinberg equilibrium  $P>5x10^{-5}$  in controls. <sup>28,129,130,132</sup> Each cohort then underwent imputation (Impute 2.2) to a shared reference panel from the 1000 Genomes project (Phase I integrated variant set release (v3), March 2012, NCBI build 37(hg19 Following imputation SNPs were quality controlled by quality score (Q>0.9) and minor allele frequency (MAF>0.05 in controls) filtering, resulting in a common set of 5331120 SNPs across all discovery phase participants.

The metaGWAS analysis was conducted using the METAL software package  $^{133}$  on the BCISNPmax database platform (version 3.5, BCI Platforms, Espoo, Finland). METAL was implemented using the sample size scheme with weighting for each cohort being two times the case number. The analysis was adjusted for genomic inflation ( $\lambda$ ) in each cohort.

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## Glioma GWAS

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1057	The aggressive periodontitis control sample consists of three independent studies:
1058	1. The Heinz-Nixdorff-Recall (HNR) was described in Schmermund, A., et al. Assessment of
1059	clinically silent atherosclerotic disease and established and novel risk factors for predicting

1060	myocardial infarction and cardiac death in healthy middle-aged subjects: rationale and design of the
1061	Heinz Nixdorf RECALL Study. Risk Factors, Evaluation of Coronary Calcium and Lifestyle. Am
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1075	2. The Dortmund Health Study (DOGS) is described in Berger, K. et. al. DHS: The Dortmund
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1083	DOGS is represented by Klaus Berger <sup>1</sup> and Jürgen Wellmann <sup>1</sup>
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3. The FOCUS (Food chain plus) control sample is described in Muller, N., *et al.* IL-6 blockade by monoclonal antibodies inhibits apolipoprotein (a) expression and lipoprotein (a) synthesis in humans. *J Lipid Res* **56**, 1034-42 (2015). FOCUS was supported by the Federal Ministry of Education and Research BMBF (FKZ 0315540A). FOCUS is represented by Matthias Laudes<sup>1</sup> Clinic of Internal Medicine I, University Medical Center Schleswig-Holstein, Kiel, Germany

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## The International Genomics of Alzheimer's Project (IGAP)

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Material and methods

1111	International Genomics of Alzheimer's Project (IGAP) is a large two-stage study based upon
1112	genome-wide association studies (GWAS) on individuals of European ancestry. In stage 1, IGAP
1113	used genotyped and imputed data on 7,055,881 single nucleotide polymorphisms (SNPs) to meta-
1114	analyse four previously-published GWAS datasets consisting of 17,008 Alzheimer's disease cases
1115	and 37,154 controls (The European Alzheimer's disease Initiative – EADI the Alzheimer Disease
1116	Genetics Consortium - ADGC The Cohorts for Heart and Aging Research in Genomic
1117	Epidemiology consortium - CHARGE The Genetic and Environmental Risk in AD consortium -
1118	GERAD). In stage 2, 11,632 SNPs were genotyped and tested for association in an independent set
1119	of 8,572 Alzheimer's disease cases and 11,312 controls. Finally, a meta-analysis was performed
1120	combining results from stages 1 & 2.
1121	
1122	The Japanese Collaboration Team for GWAS of Panic Disorder
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