# Imperial College London 

National Heart and Lung Institute<br>Department of Respiratory Epidemiology and Public Health

# The use of dietary patterns empirically derived from Principal Components Analysis and alternative strategies to identify associations between diet and disease. 

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A thesis submitted for the degree of
Doctor of Philosophy of Imperial College London

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#### Abstract

Dietary patterns derived empirically using principal components analysis (PCA) are widely employed for investigating diet-disease relationships. The aim of the study was to investigate whether PCA performed better at identifying associations between diet and disease than analysing each food on the FFQ separately, a process we refer to as exhaustive single food analysis (ESFA).

A systematic review of nutritional epidemiology literature relating to the use of PCA in identifying dietary patterns in observational and cohort studies from 2004-2009 was employed. Furthermore, we simulated diet and disease data using real food frequency questionnaire data and assuming that a number of foods or dietary pattern intakes were causally associated with disease. In each simulation, ESFA and PCA were employed to identify foods associated with disease using logistic regression, allowing for multiple testing and adjusting for energy intake. ESFA was further adjusted for principal components, foods which were significant in unadjusted ESFA, and propensity scores. For each method, we investigated the power, with which we could identify an association between diet and disease, and the power and false discovery rate (FDR) for identifying associations with specific food intakes. We apply our innovative methodology to a real dietary dataset (GA ${ }^{2}$ LEN survey).

ESFA had greater power to detect an association of diet with disease than PCA, and greater power and lower FDR for identifying associations with specific foods. FDR increased with increasing sample size using both methods. However, when ESFA was adjusted for foods that were significant in unadjusted ESFA, FDRs were controlled successfully at the desired level of $20 \%$.

Our results raise questions about the use of PCA in nutritional epidemiology. Adjusted ESFA identifies foods that are causally linked to disease with a low rate of false discoveries, and surprisingly good power. These findings were not fully supported from the analysis of the $\mathrm{GA}^{2}$ LEN data-set.


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## 1 Introduction

1.1 Dietary patterns and nutritional epidemiology
1.2 Premises
1.3 Data sources
1.4 Outline of thesis
"In the Pre-Socratic Era, mathematical knowledge was thought to supply an ideal, from which every-day empirical knowledge fell short. Specifically, if the world of observation doesn't fit the world of mathematics so much the worse for world of observation."

Bertrand Russell

### 1.1 Dietary pattern analysis and nutritional epidemiology

Epidemiology has been defined as the study of the patterns of disease occurrence in populations and of the factors that influence those patterns (Rothman \& Greenland, 1998). Nutritional epidemiology is therefore the study of the nutritional determinants of disease and is concerned with effects of diet on chronic diseases. These effects are multifactorial in origin and may take years, if not decades, to develop (Willett, 1998).

The public have an increasing interest in the role of diet in the aetiology of certain chronic diseases and has been often confused by the contradictory results of the empirical findings of nutritional epidemiology.

These contradictory results in observational studies are due to the complex nature of diet and the vast number of potentially relevant nutrients and foods. Because the prevailing method of analysis has been to study the relation of outcomes to intake of single foods or nutrients, this complexity of diet may be difficult to take into account.

An alternative and increasingly popular approach, which claims to resolve the limitations of single food analysis, is to explore associations by identifying dietary "patterns" from Food Frequency Data with the use of Principal Component Analysis (PCA) (Schwerin et al., 1981, Kant, 2004, Newby \& Tucker, 2004, Hu, 2002). PCA groups food items according to the degree that they are correlated with each other, and aggregates them to distinct dietary patterns. The concept of PCA lies in the explanation of dietary behaviours of the population and the way that foods are consumed in combination with each other. However, PCA as a method applied in nutritional epidemiology raises conceptual issues and statistical problems which call for a subjective judgment by the researcher (Newby \& Tucker, 2004).

Slattery (2008) claimed that eating patterns derived from PCA characterized the diet associated disease risk better than anyone food or nutrient. However, in order for this statement to be true in disease prevention we need PCA to identify all of the foods (and only those foods) which, in combination, increase or decrease the risk of disease, and this food combination has to be easily translatable into an intervention.

### 1.2 Premises

The primary premises of this research are that

1. There is no comprehensive critical evaluation that establishes the methodological superiority of PCA despite its wide application in nutritional epidemiological studies.
2. Simpler methods than PCA could be equally or more effective for detecting diet and disease associations.

### 1.3 Aims and objectives

We aim

1. To give a historical overview of nutritional epidemiological methods for detecting dietdisease associations with a particular focus on the application of PCA.
2. To undertake a systematic review of PCA as employed in dietary pattern analysis for observational studies.
3. To compare the results of analysing each individual food on the Food Frequency Questionnaires (FFQ) separately in relation to disease risk, a process we refer to as an exhaustive single food analysis (ESFA), with empirically derived dietary patterns with the use of Principal Component Analysis (PCA) to identify diet-disease associations in diverse populations and comparing the performance of these two methods. For this we used Monte Carlo Simulations of a dietary data-set with a realistic correlation structure.
4. To compare the use of three different methods of adjustment to cope with confounding in an unadjusted EFSA.
5. To apply our innovative methodology to a real dietary dataset derived from the GA ${ }^{2}$ LEN survey and follow-up.

### 1.4 Data sources

Our two real dietary datasets and source of the food correlation matrices for our simulations were comprised of 856 adults aged 16-50 years old living in Greenwich and questioned as part of the F.L.A.G survey (Shaheen et al., 2001) and from 200 adults 29-54 years old living in Ipswich and Norwich interviewed as part of the UK ECRHS II diet survey (Hooper et al., 2010). For more detailed information see paragraph 4.8.

Our real dataset for the application of our conclusions from our Monte Carlo simulation results was comprised of 3057 adults aged 21-75 years old living in 17 centres across Europe; Belgium (Ghent), Denmark (Odense), Poland (Lodz, Katowice), Germany (Berlin, Duisburg), Portugal (Coimbra), Italy (Palermo), Sweden (Gothenburg, Stockholm, Umea, Uppsala), Netherlands (Amsterdam), FYROM (Skopje), United Kingdom (Southampton, London) and Finland (Helsinki) as part of the Global Allergy and Asthma Network of Excellence (GA ${ }^{2}$ LEN) survey. For more detailed information see paragraph 6.2.

### 1.5 Outline of thesis

- Introduction. A description of the problem studied, of the primary premises, aims and objectives is presented.
- Background. A literature review of the different methods used for dietary patterns analysis as applied in observational studies of diet with an emphasis on PCA.
- Systematic Review. A systematic review of nutritional epidemiology literature relating to the use of PCA in identifying dietary patterns in observational studies from 20042009.
- Methods. A detailed description of our Monte Carlo simulations, where we created a hypothetical population in which we tested which one of the two approaches, ESFA or PCA had greater power and lower false discovery rate, and which method of adjustment was more appropriate for dealing with confounding in ESFA.
- Results. Average estimates of percentages of power and false discovery rates for different sample sizes and different number of principal components. Average percentages of false discovery rates for different ways of adjustment of ESFA method for different sample sizes are presented.
- Analysis of GALEN data-set: effect of diet on asthma. This section presents

1. An innovative two-step analysis with the use of generalized linear models. As a first step we explore associations of each individual food in the FFQ with respiratory and allergic outcomes. As a second step, we re-run the same analysis additionally adjusting for foods that were identified as statistically significant at the first step controlling the false discovery rate at $20 \%$.
2. Dietary patterns analysis with the use of Principal Component Analysis (PCA).

All of the statistical models were adjusted for potential demographic and environmental risk factors and multiple testing. Where necessary, multilevel and meta-analytical techniques were employed within this framework of analysis to account for between-centre and within-centre variation.

- Discussion and Conclusion. Interpretation of the study results and limitations of the study are presented along with ideas for future research.


## 2 Background

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### 2.1 From single food and nutrient analysis to dietary patterns

### 2.1.1 A historical overview of the complexity of diet

Nutritional epidemiology is a scientific field with a history of over 200 years. Lind (1753) conducted one of the first clinical trials in 1776 to test the effect of citrus fruit in preventing scurvy and concluded that lemons and oranges have a protective effect on scurvy (Carpenter, 1986, Willet, 1998). In the late nineteenth century, during the era of industrialisation it was hypothesised that the inclusion of milk and vegetables in the human diet could eliminate beriberi and pellagra (Messina et al., 2001). Later on, it was proved that world epidemics of these diseases were due to vitamin and nutrient deficiency; vitamin C for scurvy, thiamine for beriberi and nicotinic acid for pellagra (Jacobs \& Steffen, 2003). At the start of the twentieth century, this fundamental evidence and the discovery of most of the vitamins and minerals led to the belief that nutrient deficiency was the primary cause of disease symptoms (Jacobs \& Steffen, 2003). So, the attention of nutritional science moved towards the investigation of specific foods and nutrients (vitamins, lipids, amino acids) with protective effects; the attention of industry to the production of nutrient supplements; and the public health policy towards prevention strategies (Mertz, 1984).

However, in the second half of the twentieth century and with the advent of westernized chronic diseases, nutrient deficiency has not been the only explanation for diet-disease associations. Although single food and nutrient analysis has continued to make important steps forward in identifying nutrient-disease associations (Shekelle et al., 1981, Knert et al., 2004, Giovannucci et al., 2006), recent results have raised questions about the aetiology of diet in certain chronic diseases. Specifically, a lack of empirical evidence from randomised controlled trials based on observational findings, has been observed in chronic diseases like asthma (Shaheen et al., 2007, Pearson et al., 2004, Fogarty et al., 2003), cancer (Greenberg et al., 1994, Schatzkin et al., 2000) and cardiovascular disease (Hennekens et al., 1996). This may be due to the moment of intervention in life, duration of the intervention, follow-up period of the intervention as well to the various conceptual and methodological limitations of single food and nutrient analysis in taking into account the complex biological and behavioural effects of diet on disease.

### 2.1.2 Limitations of single food and nutrient analysis and the dietary pattern analysis solution

Occurrence of chronic diseases in the second half of the twentieth century has been linked with multiple factors and not only with vitamin or nutrient deficiency. As Willet pointed out in the second chapter of his book Nutritional epidemiology major energy sources (proteins, carbohydrates, fats, alcohol), food additives, chemical contaminants (pesticides, herbicides), microbial toxic contaminants, inorganic contaminants, chemicals formed in the cooking or processing of food and natural toxins have played a role in understanding the relation of diet to disease. Moreover, dietary causes of disease have posed a significant challenge to the science of epidemiology, since the diet of an individual represents a large and complex set of exposures that are difficult to measure (Willet, 1998).

Independent effects of a dietary component of interest may be difficult to identify and can be partly confounded by other dietary exposures and an individual's socio-demographic (GexFabry, Raymond \& Jeanneret, 1988, Northstone, Emmett \& Rogers, 2008) and behavioural patterns. In addition, collinear associations between food or nutrient variables can increase the uncertainty of the estimated models (Jacques \& Tucker, 2001, Michels \& Schulze, 2005, Randall et al., 1990, Kant et al., 1991). Individuals who try to eat a healthy diet are likely to lead a healthy lifestyle in general and it is not always possible to measure all important markers of a healthy lifestyle (Michels, 2003). A good example for potential confounding and lifestyle factors has been given in relation to per capita meat intake and colon cancer (Armstrong \& Dol, 1975). However, in rich countries, people can afford to eat fat rather than starchier grain products. Some aspects of the diet in these countries, or other factors in the life-style, probably do cause certain kinds of cancer and protect against other kinds. So far, epidemiologists can identify only a few of these factors with any real confidence and fat is not among them (Willet et al. 1998).

Furthermore, there are unmeasured additive and interactive effects when foods are consumed in combination, which are difficult to take into account by single food or nutrient analysis (Sacks et al., 1995, Newby et al., 2006b). Studies have claimed that dietary constituents may interact with each other biologically in complex ways, for instance one nutrient may modify the absorption, metabolism or requirement for another nutrient (Willet, 1998, Sacks et al., 1995) and these interactions may have an effect on health. For example, vitamin D acts
synergistically with interleukins to inhibit the proliferation of MC-7 breast cancer cells and that calcium supplementation decreases the risk of adenoma but only in patients who consume a low fat diet (Messina et al., 2001).

Finally, since there is a large number of foods that are consumed by any population, when we try to identify associations between food or nutrient intake and disease, we are testing a large number of hypotheses and indiscriminate multiple testing can result in chance findings (Teo \& Chong, 2006, Benjamini \& Hochberg, 1995).

Over the last thirty years, based on the hypothesis that the complexity of diet-disease associations could not adequately be addressed using only traditional approaches, there has been an explosion in the use of dietary pattern analysis in nutritional epidemiological studies (Kant, 2004, Newby \& Tucker, 2004, Hu, 2002). The main idea behind the use of dietary patterns is that diet of an individual may usefully be described in terms of a limited number of continuous variables, each representing a selection of different foods either eaten together of mutually excluded from the diet. So, since traditional approaches could not always take into account the complexity of the diet, and the use of a large number of different food intakes creates problems of multicollinearity and multiple testing, an alternative is to look at a small number of dietary dimensions each made up of a combination of foods. This alternative method was supported by two randomised controlled clinical trials where a dietary pattern approach seemed to be effective in lowering blood pressure (Dietary Approaches to Stop Hypertension (DASH) trial) and be protective in the recurrence of a number of potential outcomes after a first myocardial infarction (Lyon Heart Study).

In the first trial, 459 adults were randomly assigned and received a control diet, a diet that was high in fruits and vegetables and a combination diet that was high in fruits, vegetables and low-fat dairy products with reduced saturated and total fat for eight weeks. There was a decrease in the systolic and diastolic blood pressure of the individuals for the combination diet ( 5.5 and 3.0 mm Hg reduction of systolic and diastolic blood pressure), and for the diet rich in fruits and vegetables ( 2.8 and 1.1 mm Hg reduction of systolic and diastolic blood pressure) compared to the people who received a control diet (Appel et al., 1997).

In the second trial, a Mediterranean type diet (patients assigned to the experimental group were asked to comply with a Mediterranean-type diet) compared to a prudent Western-type diet (patients of the control group received no dietary advice from the investigators but
nonetheless were advised to follow a prudent diet by their attending physicians) reduced the number of adverse events of 423 patients after their first myocardial infarction (de Lorgeril et al., 1999). Specifically, all-cause and cardiovascular mortality and the combination of recurrent myocardial infarction and cardiac death (RR: $0.28 ; 95 \% \mathrm{CI}: 0.15-0.53$ ), or the preceding plus major secondary end points (unstable angina, stroke, heart failure, pulmonary or peripheral embolism) (RR: $0.33 ; 95 \%$ CI:0.21-0.52), or the preceding plus minor events requiring hospital admission (RR:0.52; 95\% CI:0.38-0.74) were reduced in the Mediterranean type diet group compared to the prudent Western-type diet group.

Dietary patterns (also referred to as food or eating patterns) can be derived either a priori or empirically from a set of data.

### 2.2 A priori dietary patterns

A priori dietary patterns use current nutritional knowledge from empirical research or theory based on prevailing hypotheses and guidance about the role of food items and nutrients in disease prevention. Diet is assessed by a dietary index that a research group has created in order to rank the presence or absence of certain food or nutrient characteristics, and the resulting score is used as an overall measure of dietary quality. As reported in a review by Kant et al. 2004 (Kant, 2004) the reported dietary indexes/scores as dietary patterns can be grouped into three major categories:
i) Dietary variety-based scores, such as the Dietary variety score based on a cumulative number of food items consumed on 15 consecutive days (Drewnowski et al., 1997), or the CARDIA dietary questionnaire in which diet variety was defined as the number of unique food items reported (Slattery et al., 1997).
ii) Scores derived from food related dietary guidance, such as the Diet Quality Index (DQI) (Haines et al., 1999) and the Health Eating Index (HEI) (Hann et al., 2001) which are analytic scoring tools (in a scale of 100 points) being used to measure compliance with dietary recommendations and guidelines.
iii) Mediterranean dietary scores which assess the conformity to the traditional Mediterranean diet with a 10 unit scale which relies on nine dietary components that capture the essence of the traditional Mediterranean diet. In the Mediterranean dietary score food items such as vegetables, legumes, fruits and nuts, fish and seafood and
cereals are presumed to be beneficial for health, whereas meat and dairy products are presumed to be harmful (Trichopoulou et al., 2003, Chatzi et al., 2007).

### 2.3 Empirically derived dietary patterns

Empirically derived eating patterns are not defined a priori but are data driven. Statistical methods are used to generate patterns from collected dietary data. Dietary assessment methods that are commonly used to collect the dietary data are food frequency questionnaires (FFQ), diet recalls or dietary records. These three dietary assessment methods are different in terms of how costly they are and what measurements of diet they provide (Willet, 1998). In brief, FFQ's are usually measure habitual consumption (portions of a food item per day, or units per week) of an individual over a long period (usually one year), while diet recall's and dietary records collect data over a small period (usually one week or less). In nutritional epidemiology datadriven methods that are commonly used to derive dietary patterns are presented below with a detailed description on Principal Component Analysis.

### 2.3.1 Cluster Analysis

Cluster analysis is the second most popular method after Principal Component Analysis for identifying dietary patterns of the population (Kant, 2004, Newby \& Tucker, 2004, Hu, 2002). The basic aim of cluster analysis in nutritional studies is to find natural groupings, if any, of a set of individuals according to their dietary intake of specific food items. Newby et al.'s 2004 review indicated 2 commonly reported methods used in nutritional epidemiology for cluster analysis, which are going to be described in brief; Ward's and K-means (Chatfield \& Collins, 1980, Anderson, 2003). Studies usually used both approaches in order to decide on the number of clusters (dietary patterns) that should be derived.

Ward's method is a hierarchical clustering method and is designed to optimize the minimum variance of individual's dietary intakes within clusters, which represent our dietary patterns. It uses analysis of variance at each merging step, considering all possible pairs of clusters and retaining the one with the smallest increase in the error sum of squares (Ward, 1963).

The K-means algorithm is a non-hierarchical simple, iterative procedure which is based on the definition of a point (centroid) in the space of records which represents an average location of the particular cluster (using the squared Euclidian distances between observations to determine cluster position). Thus, the coordinates of this point are averages of dietary intakes of all subjects who belong to the cluster. Each individual is positioned in space on the basis of
intake of numerous foods. Food choices common to all contribute less to cluster formation than those choices made by some and not by others. Clusters are named according to food items that on average contributed relatively more to total energy intake. The reason for this standardization was to account for differences in total energy needs due to demographic and lifestyle factors (Anderson et al., 2011). Labelling of the clusters could also be performed on the basis of the percentage of people in each cluster consuming lower or higher-than-median value of the food items stratified for.

After deciding on the number of clusters and the number of foods that constitute them, diet and disease associations are examined with clusters being the categorical exposure variables (Kant, 2004, Reedy et al., 2010, Engeset et al., 2005, Newby, Muller \& Tucker, 2004b, Costacou et al., 2003).

### 2.3.2 Reduced rank regression

Reduced Rank Regression (RRR) applied in nutritional epidemiology is described in detail in Hoffman et al.'s 2004 paper (Hoffmann et al., 2004). In brief, RRR is a statistical data reduction technique which defines linear combinations of food intakes that maximally explain intermediate markers of disease. RRR requires two sets of variables for the identification of dietary patterns; predictors which could be the dietary intakes derived from a food frequency questionnaire and responses which could be nutrient intakes (Schulz et al., 2005, Hoffmann et al., 2005, Nothlings et al., 2008) or biomarkers that are in the pathways between the foods and health outcomes (McNaughton, Mishra \& Brunner, 2008, Liu et al., 2009). RRR determines linear functions of predictors by explaining as much variation in a group of response variables. Finally, factor scores of the predictors are used as the exposure variables and their association with disease is investigated (Kroke, 2004).

### 2.3.3 Other data-driven methods

In brief, other less commonly data driven methods applied in food frequency data are:
i) Conditional Gaussian mixture modelling. A latent variable solution is proposed for dietary pattern analysis by using a finite mixture model to identify mutually exclusive subgroups of individuals with different dietary profiles. The main assumption of this method is that if subgroups exist in a sample of individuals who are distinguished by their dietary profiles, these subgroups would be expected to have different food intake
probability distributions which could be explained by a conditional Gaussian mixture model (Fahey et al., 2007).
ii) Cluster analysis of principal component scores (see a detailed description of principal component scores at paragraph 2.4.5) instead of food items. In Brief, principal component scores are calculated by applying principal components analysis on a Food Frequency Questionnaire, and then these obtained principal component scores entered in the cluster analysis procedure (He et al., 2009).
iii) Treelet transform (TT) which can be viewed as an amalgamation of PCA and hierarchical clustering methods (Gorst-Rasmussen et al., 2011). TT algorithm locates the two variables in the dataset with the largest correlation, performs PCA on them and creates a corresponding score. A merge is indicated in the cluster tree. This scheme continues until all of the variables have joined the cluster tree. TT procedure results in the production of pattern scores by aggregating dietary intake values according to correlation. Furthermore, TT singles out a smaller number of interrelated dietary variables than PCA by introducing sparsity to the principal component loadings; that is making a many loadings exactly zero (for a detailed description of principal component loadings see paragraph 2.4.7).
iv) In some studies researchers have employed factor analysis. Although factor analysis shares aims with PCA, it is not recommended for the analysis of nutritional data. The reason is that PCA is commonly used to define dietary patterns because the principal components are certain mathematical functions of the observed variables, whereas common factors are not expressible by the combination of the observed variables. Even when people say that they are employing factor analysis in nutritional studies they may be employing PCA, and this misconception depends on the statistical package that these studies have used (Agur-Collins et al., 2009, Hughes et al. 2009, Yuna et al., 2009).

### 2.4 Principal Component Analysis (PCA) and nutritional epidemiology

### 2.4.1 A brief history of PCA and its application in nutritional epidemiology

PCA is one of the oldest techniques of multivariate analysis. It was introduced by Pearson (1901) and developed by Hotelling in 1933. Over the last 80 years, as Jolliffe points out there has been a wide application of the Principal Component Analysis (PCA) method in the fields of psychology, agriculture, genetics, biology, chemistry, physics, meteorology quality control (Jolliffe, 2010). The use of PCA in nutritional epidemiology as a method for identifying dietary patterns is the focus of this thesis.

PCA as a method for reduction of measurements used to assess nutritional status dates back to a study by Drion published in 1961. In this study fourteen nutrients entered the PCA, and four principal components were identified which expressed (i) level of total consumption, (ii) relative importance of animal versus vegetable protein, (iii) quantity of fruits and vegetables (other than potatoes) consumed and (iv) the quantity of butter and margarine consumed (Drion, 1961). Another study in the 1970's employed PCA to reflect underlying processes that had created the correlation among thirty-two variables (demographic, socio-economic, anthropometric, dietary, biochemical and urinary) which were used to assess nutritional status (Gurthie et al., 1973). It wasn't until 1981 when the association between 7 eating patterns derived from PCA and health was examined from the TEN-STATE and HANES I survey (Schwerin et al., 1981). Although the PCA approach dates back half a century, there has been an explosion in the use of the method in recent years.

Specifically, our systematic review (see Chapter 3) identified 163 papers employing PCA in observational studies between 2004-2009 and other systematic reviews identified 41 papers between 1998-2004, and 13 between 1981-1997 (Kant, 2004, Newby \& Tucker, 2004). The main reason why the PCA method applied in nutritional epidemiology gained so much attention after 1998 was a paper published in the American Journal of Epidemiology by Slattery et al., which associated a "Western" dietary pattern with increased risk of colon cancer and identified a "prudent" dietary pattern as protective against colon cancer in a population-based study conducted in Northern California, Utah, and Minnesota (Slattery et al. 1998). These two patterns of diet identified by Slattery were also identified in two other large US cohorts, where a "prudent" pattern characterised by intake of vegetables, fruit, legumes, whole grain, fish and poultry, and a "Western" pattern characterised by intake of red meat, processed meat, refined grains, sweets and dessert, French fries, and high-fat dairy products
were identified and subsequently linked to coronary heart disease (CHD) (Hu et al., 2000, Fung et al., 2001), colon cancer (Wu et al., 2004, Fung et al., 2003) and chronic obstructive pulmonary disease (Varraso et al., 2007a, Varraso et al., 2007b). This accumulating evidence along with the methodological problems that PCA claims to solve boosted the application of the method in nutritional studies (see Chapter 3; Figure 3.2.1). However, before we provide the methodological reasons why PCA became such a popular method in nutritional epidemiology, we will give a more detailed description of the method.

### 2.4.2 Description and general purpose of PCA in nutritional epidemiology

PCA, as applied in nutritional epidemiology, is a multivariate statistical method which uses food frequency questionnaire data (Willett, 1998, Hu, 2002), dietary records (Perrin et al., 2005) or dietary history questionnaire (Yannakoulia et al., 2008, Robinson et al., 2009) to aggregate information, reflect underlying processes and explain the variance-covariance structure of a set of correlated food intake variables. Specifically, PCA examines the relationships among a set of k correlated dietary exposures by transforming them to a new smaller set of uncorrelated overall dimensions of diet, while aiming to explain as much of the variation present in the original dietary exposures; these overall dimensions of diet are called principal components, or in our case, dietary patterns.

Although k components (dietary patterns) are required to reproduce the total system variability of the k original food intake variables, often much of this variability could be accounted from a smaller number of p components (dietary patterns). So, the original dataset of n measurements of k food intake variables could be reduced to a data set of n measurements of p components. In the next paragraph we will try to describe PCA more formally.

### 2.4.3 Derivation of population dietary patterns (principal components)

Let's assume that we have a random vector of $k$ correlated food intake variables
$X^{T}=\left[X_{l}, X_{2} \ldots, X_{k}\right]$ that have a covariance matrix $k_{x} k$

$$
V\left[X^{\mathrm{T}}\right]=\Sigma=\left[\begin{array}{cccc}
\sigma_{11} & \sigma_{21} & & \sigma_{k 1}  \tag{2.4.3.1}\\
\sigma_{12} & \sigma_{22} & & \sigma_{k 2} \\
\vdots & & & \\
\sigma_{1 k} & \sigma_{2 k} & \ldots & \sigma_{k k}
\end{array}\right]
$$

PCA solves the problem of finding a new set of $k$ variables $Y^{T}=\left[Y_{1}, Y_{2}, . ., Y_{k}\right]$ called dietary patterns (principal components) which are uncorrelated linear combinations of the original food intake variables $X_{1}, X_{2}, \ldots, X_{k}$ and whose variances are large as possible and decrease from first to last. The first principal component is the linear combination of food intakes with maximum variance. The second principal component is the linear combination of food intakes with maximum variance subject to being uncorrelated with the first component. The $k^{t h}$ principal component is the linear combination of food intakes with maximum variance subject to being uncorrelated with all of the previous components.

More formally, we consider the linear combinations as
$Y_{1}=a_{1}^{\mathrm{T}} X=a_{11} X_{1}+a_{21} X_{2}+\ldots+a_{k 1} X_{k}$
$Y_{2}=a_{2}^{\mathrm{T}} X=a_{12} X_{1}+a_{22} X_{2}+\ldots+a_{k 2} X_{k}$
$Y_{k}=a_{k}^{\mathrm{T}} X=a_{1 k} X_{1}+a_{2 k} X_{2}+\ldots+a_{k k} X_{k}$
where
$a_{j}^{\mathrm{T}}=\left[a_{1 j}, \ldots, a_{k j}\right]=\left[\begin{array}{cccc}a_{11} & a_{21} & & a_{k 1} \\ a_{12} & a_{22} & & a_{k 2} \\ \vdots & & & \\ a_{1 k} & a_{2 k} & \ldots & a_{k k}\end{array}\right](j=1,2, \ldots, k)$
is a vector of constants (coefficients).
The general linear form of equation (2.4.3.2) is given by

$$
\begin{equation*}
Y=\alpha^{T} X \tag{2.4.3.4}
\end{equation*}
$$

Each dietary pattern $Y_{j}(j=1,2, \ldots, k)$ is a weighted sum of the food intake variables $X_{i}$ 's $(i=1,2, \ldots, k)$, and $a_{i j}$ 's $(i=1,2, . ., k j=1,2, \ldots, k)$ are the weights of the food intake variable $i$ on pattern j. According to Chatfield and Collins (Chatfield \& Collins, 1980; page 24) the mean of vector $Y$ as described by equation (2.4.3.4) is given by the formula
$E(Y)=a^{T} E(X)(2.4 .3 .5)$
Moreover, variance and covariance of dietary patterns $Y_{j}$ s $s(j=1,2, ., k)$ are given from the formulas
$\operatorname{Var}\left(Y_{j}\right)=a_{j}^{\mathrm{T}} \Sigma a_{j}$
$\operatorname{Cov}\left(Y_{j}, Y_{w}\right)=a_{j}^{\mathrm{T}} \Sigma a_{w}(j \# w ; j=1, \ldots, k ; w=1, \ldots, k)$

From equation (2.4.3.6), we can maximize arbitrarily the variance $\operatorname{Var}\left(Y_{j}\right)$ of dietary pattern $Y_{j}(j=1,2, . ., k)$ by multiplying the factor $a_{j}^{\mathrm{T}}(j=1,2, . ., k)$ with a constant factor. So in order to have maximum variance and a deterministic value, $a_{j}^{\mathrm{T}} \mathrm{s}$ should be subject to the specific constraints of algebraic orthogonality.

$$
\begin{align*}
& a_{j}^{\mathrm{T}} a_{j}=1(j=1, . ., k) \\
& \operatorname{Cov}\left(Y_{j}, Y_{w}\right)=a_{j}^{\mathrm{T}} \Sigma a_{w}=0(j \# w ; j=1, \ldots k ; w=1, \ldots, k) \tag{2.4.3.9}
\end{align*}
$$

Dietary patterns $Y_{1}, Y_{2}, \ldots, Y_{k}$ are derived in decreasing order of importance in terms of explaining the largest proportion of total variance of the dietary intakes of the population. In order to do so, we need to use a mathematical optimization procedure for finding the minima and maxima of a function of several variables subject to a specific constraint through the method of Lagrange multipliers (Chatfield \& Collins, 1980, Johnson \& Wichern, 1982). Particularly, in order to maximize the variance of a principal component (dietary pattern) we need to introduce a new variable $\lambda$, which is called a Lagrange multiplier and it is the stationary point for the Lagrange function (for this point the derivative of the Lagrange function is zero and is the point when the function stops to increase or decrease). The mathematical problem that this procedure poses could be easily solved with the help of matrix algebra. So, a non-null solution for determining the $a_{j}$ 's in order for the dietary patterns to have the required properties of orthogonality is by finding the eigenvalues $\lambda_{1} \geq \lambda_{2} \geq \ldots \geq \lambda_{k} \geq 0$ of the covariance matrix $\Sigma$ of the food intake variables which are the roots of the equation

$$
\begin{equation*}
|\Sigma-\lambda I|=0 \tag{2.4.3.10}
\end{equation*}
$$

To each eigenvalue $\lambda_{i}(i=1, \ldots, k)$ corresponds a vector $c_{i}$, which is called an eigenvector such that
$\Sigma c_{i}=\lambda_{i} c_{i}$

Principal components (dietary patterns) are the eigenvectors defined by equation (2.4.3.11).

Also using equation (2.4.3.10) and (2.4.3.6) we have that

$$
\begin{equation*}
\operatorname{Var}\left(Y_{1}\right)=a_{1}^{\mathrm{T}} \Sigma a_{1}=a_{1}^{\mathrm{T}} \lambda_{1} I a_{1}=\lambda_{1} \tag{2.4.3.12}
\end{equation*}
$$

Thus, since we want to maximize the variance $\operatorname{Var}\left(Y_{1}\right)$ of the $\mathbf{1}^{\text {st }}$ principal component (dietary pattern) subject to the constraint that $a_{1}^{\mathrm{T}} a_{1}=1$ we want to choose the largest eigenvalue $\lambda_{1}$ and the corresponding eigenvector $a_{1}^{\mathrm{T}}=\left[a_{11}, \ldots, a_{k 1}\right]$ for this eigenvalue (Anderson, 2003, Bartholomew \& Steele, 2002). In addition, in order to maximize the variance $\operatorname{Var}\left(Y_{2}\right)$ of the $2^{\text {nd }}$ principal component (dietary pattern) $Y_{2}$ subject to the following constraints
$a_{2}^{\mathrm{T}} a_{2}=1$
$\operatorname{Cov}\left(Y_{1}, Y_{2}\right)=a_{1}^{\mathrm{T}} \Sigma a_{2}=0$
we will choose $\lambda_{2}$ to be the second largest eigenvalue for the corresponding eigenvector $a_{2}^{\mathrm{T}}=\left[a_{12}, \ldots, a_{k 2}\right]$. So accordingly, in order to maximize variance $\operatorname{Var}\left(Y_{j}\right)$ of the $\boldsymbol{j}^{\text {th }}$ principal component (dietary pattern) $Y_{j}$ we will choose $\lambda_{j}$ to be the $j^{\text {th }}$ largest eigenvalue corresponding to the eigenvector $a_{j}^{\mathrm{T}}=\left[a_{1 j}, \ldots, a_{j}\right]$.

So as a general framework
$\operatorname{Var}\left(Y_{j}\right)=\lambda_{j}, j=1, \ldots, k$ (2.4.3.15).

In the case that some of the eigenvalues of the covariance matrix $\Sigma$ are equal, there is not a unique way to choose the corresponding eigenvectors but they should always be chosen to be orthogonal.

To sum up, this derivation of the Principal Components (dietary patterns), weights (coefficients) and variances as eigenvectors and eigenvalues of a covariance matrix is standard and appears in most text books. All equations presented above are taken from Chatfield \& Collins and Johnson \& Wichern (Chatfield \& Collins, 1980, Johnson \& Wichern, 1982).

Furthermore, we would like to express the total variance being explained by the principal components (dietary patterns) of the original food intake variables. So, let's assume again that $\left(k_{x} k\right)$ matrix of eigenvectors
$A=\left[\begin{array}{cccc}a_{11} & a_{21} & \ldots & a_{k 1} \\ a_{12} & a_{22} & & a_{k 2} \\ \vdots & & & \\ a_{k 1} & a_{k 2} & \ldots & a_{k k}\end{array}\right]$
And the ( $k_{x} 1$ ) vector of principal components

$$
Y=\left[\begin{array}{c}
Y_{1}  \tag{2.4.3.17}\\
Y_{2} \\
\vdots \\
Y_{k}
\end{array}\right]
$$

with the covariance matrix of Y given by

$$
\Lambda=\left[\begin{array}{cccc}
\lambda_{1} & 0 & \ldots & 0  \tag{2.4.3.18}\\
0 & \lambda_{2} & \ldots & 0 \\
\vdots & & & \\
0 & & \ldots & \lambda_{k}
\end{array}\right]
$$

From equations (2.4.3.1), (2.4.3.6) and (2.4.3.18) we can derive that

$$
\begin{equation*}
\Lambda=\operatorname{Var}(Y)=A^{\mathrm{T}} \Sigma A \tag{2.4.3.19}
\end{equation*}
$$

The spectral decomposition of a square matrix could be given by
$\Sigma=A \Lambda A^{T}$
since A is an orthogonal matrix and

$$
A A^{T}=I \quad \text { (2.4.3.21) }
$$

Furthermore in linear algebra the trace of two square $\mathrm{k}_{\mathrm{x}} \mathrm{k}$ matrices A and B is the sum of their diagonal elements

$$
\begin{equation*}
\operatorname{trace}(A)=a_{11}+a_{22}+\ldots+a_{k k}=\sum_{i=1}^{k} a_{i i} \tag{2.4.3.22}
\end{equation*}
$$

$\operatorname{trace}(B)=\beta_{11}+\beta_{22}+\ldots+\beta_{k k}=\sum_{i=1}^{k} \beta_{i i}$
with the properties
$\operatorname{trace}(A B)=\operatorname{trace}(B A) \quad(2.4 .24)$
so in our case

$$
\begin{equation*}
\operatorname{trace}(\Sigma)=\sigma_{11}+\sigma_{22}+\ldots+\sigma_{k k} \tag{2.4.3.25}
\end{equation*}
$$

$$
\begin{equation*}
\operatorname{trace}(\Lambda)=\lambda_{1}+\lambda_{2}+\ldots+\lambda_{k} \tag{2.4.3.26}
\end{equation*}
$$

So combining (2.4.24), (2.4.3.25), (2.4.3.26), (2.4.3.15), (2.4.3.21) we have that

$$
\begin{equation*}
\sum_{i=1}^{k} \operatorname{Var}\left(X_{i}\right)=\operatorname{trace}(\Sigma)=\operatorname{trace}\left(A^{T} \Lambda A\right)=\operatorname{trace}\left(A^{T} A \Lambda\right)=\operatorname{trace}(\Lambda)=\sum_{j=1}^{k} \lambda_{j}=\sum_{j=1}^{k} \operatorname{Var}\left(Y_{j}\right) \tag{2.4.3.27}
\end{equation*}
$$

So the total population variance of food intake variables $X_{i}$ 's could be expressed by the sum of the variance of the dietary patterns $Y_{j}$ 's in terms of their eigenvalues and the proportion of total variance of the food items explained by the dietary patterns is given by
$\frac{\lambda_{j}}{\lambda_{1}+\lambda_{2}+\ldots+\lambda_{k}}$
Consequently, the proportion explained by the first $p$ dietary patterns together is $\frac{\lambda_{1}+\lambda_{2}+\ldots+\lambda_{p}}{\lambda_{1}+\lambda_{2}+\ldots+\lambda_{k}}$

### 2.4.4 Principal Component Analysis with the use of a correlation matrix

In nutritional epidemiology, food intake variables $X_{i}$ 's are assumed to be continuous and usually standardised to have a mean zero and variance 1 . So, when food intake variables are standardized before entering the PCA procedure, instead of using the covariance matrix $\Sigma$ we effectively use the correlation matrix P. Let's assume that $r_{i j}$ is the correlation coefficient between food item $i$ and food item $j$, then

$$
\begin{equation*}
r_{i j}=\frac{\operatorname{Cov}\left(X_{i}, X_{j}\right)}{\operatorname{Var}\left(X_{i}\right) \operatorname{Var}\left(X_{j}\right)} i=1,2, \ldots, k, j=1, \ldots, k \tag{2.4.4.1}
\end{equation*}
$$

$$
\mathrm{P}=\left[\begin{array}{cccc}
1 & r_{12} & & r_{1 k}  \tag{2.4.4.2}\\
r_{21} & 1 & & r_{2 k} \\
\vdots & & & \\
r_{k 1} & r_{k 2} & \cdots & 1
\end{array}\right]
$$

All the properties for finding the principal components of a covariance matrix of our food intake variables are valid when we use the correlation matrix. As we can observe from the correlation matrix P all the diagonal terms are 1 . So in this case, and according to equations (2.4.3.22) (2.4.3.23) and (2.4.3.27)
$\operatorname{trace}(P)=r_{11}+r_{22}+\ldots+r_{k k}=\sum_{i=1}^{k} r_{i i}=k$
and the proportion of the total variance for $j^{\text {th }}$ component (dietary pattern) will be given by

$$
\begin{equation*}
\frac{\lambda_{j}}{k} \tag{2.4.4.3}
\end{equation*}
$$

Standardization and use of the correlation matrix needs to be applied in dietary data because of their widely different scales (e.g. bread could be measured in grams/per day and beer in pints /per day). If they are not standardized, PCA will simply pick up the variable that has the highest variance as the direction of greatest variability (Jolliffe, 2010). Another argument for using the correlation matrix is that we need to make food intake variables directly comparable between each other. This approach will also help to make principal components (dietary
patterns) more comparable across different studies and populations (Chatfield \& Collins, 1980, Jolliffe, 2010).

Although we follow the exact same mathematical procedure for deriving the principal components, eigenvalues and eigenvectors provided from the correlation matrix $P$ are different from that of the covariance matrix. Furthermore, principal components derived from the correlation matrix don't give the same information with the covariance matrix. However, variables that tend to have large weights to one component tend to be highly correlated with that component (Johnson \& Wichern, 1982, Jolliffe, 2010). Finally, there are appropriate methods discussed in detail by Chatfield \& Collins and Jolliffe when eigenvalues of the covariance or correlation matrix are small, equal or zero (Chatfield \& Collins, 1980; page 62, Jolliffe, 2010; page 27).

### 2.4.5 Derivation of sample dietary patterns (principal components)

In nutritional epidemiology, we don't usually have information about the population under investigation but for a specific sample. So, using similar notations with that of paragraphs 2.4.3 and 2.4.4, we can assume that we have n independent drawings $x_{1}, x_{2}, \ldots, x_{n}$ from some k dimensional population with covariance matrix $S$, correlation matrix $R$, eigenvalues $\tilde{\lambda}_{1} \geq \tilde{\lambda}_{2} \geq \ldots \geq \tilde{\lambda}_{k} \geq 0$, vector of constants $\tilde{a}_{j}^{\mathrm{T}}=\left[\tilde{a}_{1 j}, \ldots, \tilde{a}_{k j}\right]$ and $p$ empirically derived principal components $y^{T}=\left[y_{1}, y_{2}, \ldots, y_{p}\right]$.

PCA is a linear procedure and was originally developed for the normal multivariate distribution and samples from it, so the variables will need to be continuous or approximately normal in order for the method to provide accurate estimates. However is a very narrow view to apply PCA only when data are approximately normal (Jolliffe, 2010). As nutritional data are highly skewed no assumptions could be satisfied for the underlying population and without these assumptions is very difficult to derive the sampling properties of the above estimates. Hence, PCA should be used in nutritional epidemiology with caution and mainly as a descriptive and not as an inferential tool.

### 2.4.6 The identification of important dietary patterns (How many dietary patterns?)

In this paragraph we present a number of rules on how to derive an appropriate number of dietary patterns. Continuing from the previous paragraph, after calculating the eigenvalues of a population or sample correlation matrix, the next step is to determine which dietary patterns are accounting for a large proportion of total variance of the standardized food items; or in other words which eigenvalues are "large" and which are "small". There are three prevailing criteria in the majority of PCA studies to determine the number of dietary patterns.

Firstly, the number of patterns could be determined by examining the percentage of total variance that the dietary patterns explain in the original dataset and is derived from the equation (2.4.4.3). Most statistical textbooks of PCA provide a rule of thumb, which determines an acceptable number of principal components when they explain around 70-90\% of total variance of the original variables (Chatfield \& Collins, 1980, Bartholomew \& Steele, 2002, Anderson, 2003, Jolliffe, 2010). However, in nutritional studies, the percentage of total variance reported is rather smaller with commonly reported percentages for dietary patterns being between $10 \%$ and $30 \%$ (median: $24 \%$, IQR: 19.9-31.3; see paragraph 3.2.5 and the review by Newby (2004)).

Secondly, as we mentioned above, it is more appropriate for nutritional data to derive dietary patterns from the correlation matrix of food intake variables. In this case, another rule of thumb is to retain those components with eigenvalues $>1$. This rule is known as the Kaiser criterion (Kaiser, 1960) and the main idea behind it is that components with eigenvalues (variances) below one explain less variation than any one of the original variables. However, in nutritional epidemiology, a number of studies who decided the number of patterns based on the cut-off point for eigenvalues used a different cut-off point (median value:1.6, IQR:1.25-2) in order to retain a small and more easily interpretable number of dietary patterns (See paragraph 3.2.5 and the review by Newby (2004)).

Thirdly, a more objective way on deciding on the number of components is with the scree plot which was named by Catell (1966) and inspired by the scree at the bottom of a mountain slope. With the eigenvalues ordered from largest to smallest, this is a plot of the magnitude of an eigenvalue versus the dietary pattern number. The aim is to identify an elbow which
corresponds to the point after which the addition of more dietary patterns explains relatively little more of the variance.

A further criterion for deciding on the number which is described in more detail by Jolliffe (2010) is the Bartlett test which tests the hypothesis that all the eigenvalues are equal, assuming that our data are following a multivariate normal distribution

### 2.4.7 Principal component interpretability

Another way to decide on the number of dietary patterns could be done on the basis that the groupings of the food items suggested by the pattern have a realistic interpretation. Interpretation of the dietary pattern could be aided with the use of component loadings, correlation coefficients and the appropriate method of rotation (see below).
2.4.7.1 Component loadings/correlation coefficients and their contribution to the label of a dietary pattern and its interpretability

Component loading is the numeric size of a food intake variable within a principal component (dietary pattern). When the covariance matrix of food items is analyzed, component loadings are rescaled coefficients $a_{i j}$ 's which we can calculate them as
$a_{i j}^{*}=\lambda_{j}^{1 / 2} a_{i j} \quad(\mathrm{i}=1, . ., \mathrm{k} ; \mathrm{j}=1, \ldots, \mathrm{k})(2.4 .7 .1 .1)$

The basic idea is to rescale the coefficients, in order that those coefficients for the most important components are larger than those of the less important components. In nutritional data, where the correlation matrix of the food items is analyzed, $a_{i j}^{*}$ may be interpreted as the correlation coefficient (or component correlation) between food item $i$ and dietary pattern $j$ (a detailed proof is provided by Chatfield \& Collins at page 63).

A positive correlation coefficient means that the food item is positively associated with the dietary pattern whereas a negative correlation coefficient reflects an inverse association with the dietary pattern. Correlation coefficients values of $\geq 0.30$ and $<-0.30$ between food items and the dietary pattern usually determine the label of the pattern (median value: 0.3, IQR: 0.30.4 ; see section 3.2.5 and review from Newby \& Tucker (2004)). The rationale for ignoring near-zero correlation coefficients is that these will correspond to only minor displacements in the direction of the variables which they multiply, so can be safely disregarded (Cadima \& Jolliffe, 1995).

### 2.4.7.2 Method of rotation

Interpretability of the patterns may increase with the appropriate method of rotation. Let's assume that we derive $p$ components from $k$ food items. By rotating the $p$ components we could find a new set of components in the p-dimensional space which are more easily interpretable. There are two families of rotations, the orthogonal and the oblique. Orthogonal family rotates the derived $p$ principal components in a way that the $p$ components or component loadings maintain the orthogonality between them and are uncorrelated. Oblique families relax the orthogonality constraint in order to gain a simpler structure for the components. Although, Catell (1978) and Richman (1986) give non-exhaustive lists of 11 and 19 such criteria, there have been two methods of rotation, varimax (orthogonal) and promax(oblique) being used in nutritional epidemiology (See paragraph 3.2.5 and the review by Newby (2004)).

Varimax rotation is being used in the majority of the studies (see section 3.2.5 and review form Newby \& Tucker (2004)) and the choice of this method over other methods of rotation is arbitrarily, that is no justification is given by the nutritional literature. This may be happening on the basis that the dietary patterns should be as uncorrelated as possible even after rotation. Fortunately, as Jolliffe (2010) points out, different choices of criteria, at least within orthogonal rotation, often make little difference to the results.

The method was suggested by Kaiser in 1958 (Kaiser, 1958) and attempts to find an orthogonal rotation that is close to a simple structure by finding $p$ principal components with few large (towards the maximum possible value) component loadings or component correlations and as many near-zero one. The total variance being explained by the components remains unchanged. So, let's assume that $R$ is a $\mathrm{k}_{\mathrm{x}} \mathrm{k}$ orthonormal rotation matrix such as

$$
R R^{T}=I(4.2 .7 .2 .1)
$$

and $R_{i j}$ is the element of the matrix in the $i^{\text {th }}$ row and $j^{t h}$ column, and $L$ is a $p_{\mathrm{x}} \mathrm{k}$ orthonormal matrix of column eigenvectors then varimax method maximizes the squared component loadings in each component according to the following definition by Stegman (2006).
$R_{\text {varimax }}=\arg \max _{R}\left(\sum_{j=1}^{k} \sum_{i=1}^{p}(L R)_{i j}^{4}-\frac{1}{p} \sum_{j=1}^{k}\left(\sum_{i=1}^{p}(L R)_{i j}^{2}\right)\right)^{2}(2.4 .7 .2 .1)$

As Jolliffe (1995) suggests, by using the normalization constraint at equation (4.2.7.2.1) the rotated varimax loadings are orthogonal but the rotated components could be correlated.

There have been 5 studies employing a promax rotation to the derived principal components in nutritional epidemiology (three of them as indicated by Newby (2004) review and two studies from our systematic review -see paragraph 3.2.5). A detailed description of the method is provided by Hendrickson and White (1964).

When the number of dietary patterns is determined and labelled then the next step is to include them as explanatory variables in a regression model.

### 2.4.8 Principal Components Regression (PCR)

Principal Components Regression (PCR) as applied in nutritional epidemiology is a regression analysis technique in which instead of using all the food intake variables as predictors, it uses the principal component scores of the derived dietary patterns which measure the conformity of an individual's diet to the given pattern.

More formally, we apply PCA in a nutritional dataset and the dietary patterns are empirically derived with the criteria mentioned in paragraphs 2.4.6 and 2.4.7. Principal component scores $\left[Y_{1}^{*}, Y_{2}^{*}, \ldots Y_{p}^{*}\right]$ are calculated for each pattern by summing the observed consumption from all standardized food items $\left(X_{1}, X_{2}, \ldots, X_{k}\right)$, weighted by the principal component loadings or correlations $a{ }^{*}{ }_{i j}$ 's according to the equation
$Y^{*}{ }_{j}=a_{1 j}^{*} X_{1}+a_{2 j}^{*} X_{2}+\ldots+a_{k j}^{*} X_{k}, j=1, \ldots, p$ (2.4.8.1)

Then, we run an ordinary least square regression with use of a generalized linear model for exploring the association between a potential outcome and the p selected derived dietary patterns.

$$
y_{j}=\beta_{o}+\beta_{1} Y_{1 j}^{*}+\beta_{2} Y_{2 j}^{*}+\ldots+\beta_{p} Y_{p j}^{*}, j=1, \ldots, k \text { (2.4.8.2) }
$$

This technique has the advantage that, in the presence of confounding and mulitcollinearity, the uncorrelated dietary pattern scores may be more easily interpretable than the correlated food intakes. Furthermore, in the regression model, with the use of the uncorrelated principal component scores instead of all the correlated food intakes the estimated regression coefficients may remain unaffected from the presence of mulitcollinearity and more stable
estimates can be obtained (Jolliffe, 2010). Because of the presence of correlation even after the rotation of a dietary pattern, usually patterns are included in the equation one at a time. Further adjustment of the dietary patterns for other socio-demographic and lifestyle confounders is usually employed.

### 2.5 Methodological justifications of Principal Component Analysis (PCA) and Principal Components Regression (PCR) and causal pathways of diet and disease at the dietary pattern, food and nutrient level.

As mentioned above (paragraph 2.1), although dietary causes of diet on disease could be logically addressed by epidemiology, the complex nature of diet has posed an unusually difficult challenge to this discipline. One of the most well studied exposures for epidemiologists is cigarette smoking and this is due to the accurate and easily obtainable quantitative information on the assessment of smoking status of an individual. On the other hand, diet represents a large set of dietary intake variables (exposures) which are highly correlated with each other. According to Willet (1998) everyone eat fat, fiber, and vitamin A and exposures cannot be characterized as present or absent; rather, they are continuous variables, often with a rather limited range of variation. As we can observe from Figure 2.5.1 (taken from page 61 (Kim \& Popkin , 2010) paper), there is numerous dietary factors which could act independently or interact with each other and potentially lead to the development of a chronic disease through different causal pathways. These pathways could be confounded by lifestyle, behavioural and early life patterns.

There are a number of methodological arguments on the rationalizing the use of PCA and PCR and how successfully they deal with the complexity of diet-disease associations (a list of them is presented for years 2004-2009 in our systematic review in Chapter 3 and Table A at the end of the thesis). Claims made for PCR include (i) that it captures additive effects of foods and nutrients on the investigated health outcome, which are too small to be detected when they are examined independently ( $\mathrm{Hu}, 2002$, Newby et al., 2006); (ii) that it captures the interactive effects of different nutrients and foods on the investigated health outcome (Hu, 2002, Varraso et al., 2007b); (iii) that it resolves aspects of confounding and mulitcollinearity between food intakes, so allowing more accurate estimates of the effect of diet on disease (Jacques \& Tucker, 2001, Randall et al., 1990); (iv) that it reduces the risk of chance findings arising from multiple statistical testing (Hu, 2002, Teo \& Chong, 2006).

Figure 2.5.1 Key pathways for diet, physical activity, and obesity on nutrition-related noncommunicable diseases (Kim \& Popkin , 2010).


The argument for (iv) is that PCA reduces the number of food items into fewer dietary patterns and consequently with PCR the number of hypotheses to be tested. However the first three arguments are highly questionable, since there is a lack of quantitative evidence to support them.

More formally, we can infer the presence of latent (i.e. unobserved) factors underlying dietary food intakes from the fact that we identify dietary patterns with the use of PCA; these identified dietary patterns are strongly correlated with a number of individual food intake variables. However, we are interested in finding causal effects of food intakes; thus, if we design a clinical trial to change food (see fruit/vegetable box in Figure 2.5.1) or nutrient (derived from foods; see saturated fat box in Figure 2.5.1) intakes of individuals, we could directly change the status of their disease risk (Figures 2.5.2, 2.5.3 and 2.5.4). Disease risk could be altered with a mechanism which could be different for each food, or these might all be foods that contain some special nutrient. (Figure 2.5.2). Associations of dietary pattern, food and nutrient intakes with disease risk could be confounded by lifestyle factors.

Figure 2.5.2 Potential pathways for diet when food intakes are not expected to be identified by PCA.


For example, observational studies have provided strong evidence that selenium is associated with the risk of asthma. Foods that are rich in selenium are Brazil nuts, sunflower seeds, fish (tuna, halibut, sardines, flounder, salmon), shellfish (oysters, mussels, shrimp, clams, scallops), meat (beef, liver, lamb, pork), poultry (chicken, turkey), eggs, mushrooms (button, crimini, shiitake), grains (wheat germ, barley, brown rice, oats) and onions. These foods are not expected to be found as foods that are highly correlated with a dietary pattern arising from PCA. However, if their standardised intakes are added together a "selenium" pattern score could be identified which is closely related to disease risk. Nevertheless, if PCA could identify this selenium pattern then we would have a different causal diagram (Figure 2.5.3). We try to address both of the causal pathways observed in the causal diagrams in figures 2.5.2 and 2.5.3 in our simulation study in Chapters 4 and 5.

Figure 2.5.3 Potential pathways for diet when food intakes are expected to be identified by PCA.


Furthermore, latent factors such as a "prudent" dietary pattern that is observed in the majority of the PCA studies of diet (see Chapter 3) could be an unobservable measure of the "desire to be healthy" which makes individuals to eat healthier, but also to exercise more and follow a prudent lifestyle. If this is true, then intervening to change food intakes will not alter disease risk of the individuals, and lifestyle factors are on the causal pathway between dietary patterns and disease (Figure 2.5.4).

Finally, there are a number of methodological considerations on the application of PCA and PCR in nutritional epidemiology. We discuss all these issues in the next paragraph.

Figure 2.5.4 Potential pathways for diet when lifestyle factors are on the causal pathway between diet and disease.


### 2.6 Methodological considerations of Principal Component Analysis (PCA) and Principal Components Regression (PCR)

To the best of our knowledge, the greater ability of PCR to capture additive effects of foods and nutrients on the investigated health outcome when compared with the analysis of single foods or nutrients has not been critically evaluated. This is one of the goals of our simulation study described in Chapters 4 and 5.

PCA, as a statistical technique and its extension PCR do not take into account statistical interactions between food or nutrient exposures, since dietary patterns are merely linear combinations of food intake variables and nutrients are linear combinations of specific food variables as well.

PCR does not solve the problem of collinearity between dietary exposures, but rather ignores the problem. If a dietary pattern is found to be associated with disease, the question still
remains as to which of the correlated or confounded food intakes that make up the dietary pattern are implicated in the association. PCR cannot be specific about which particular foods or nutrients form the dietary pattern are responsible for having a protective or positive effect with associated disease and being informative about biological relationships between dietary constituents and disease risk (McCann et al., 2001b). We are trying to address extensively this argument with our simulations.

Furthermore, in PCR, if principal components are not easily interpretable, conclusions from the regression equation may not have a clear meaning. Even when dietary patterns that are derived according to the high-variance criterion of paragraph 2.4.6 and 2.4.7 (eigenvalues $>1$, scree plot, total variance being explained by the dietary pattern) have a meaningful interpretation, the omitted dietary patterns with low variance are not necessary unimportant for the regression model (Jolliffe, 2010). In addition, as mentioned in paragraph 2.4.7, the interpretation of dietary patterns which are associated with the outcome is usually based on the correlation coefficients $\geq 0.30$ and $<-0.30$ between the food item and the derived dietary pattern, which can also suggest a subset of variables that can be used implicitly or explicitly, in a simplified interpretation (Jeffers, 1967). The rationale for ignoring near-zero loadings or correlation coefficients is that these will correspond to only minor displacements in the direction of the variables which they multiply, so can be safely discarded (Cadima \& Jolliffe, 1995). However this approach of selecting a subset of variables according to non-zero or large correlation coefficients could provide misleading results, since this approach doesn't take into account correlations between the variables (Cadima \& Jolliffe, 1995).

PCR, when one component alone contributes to the fit of the model, could fail to the account for the variation in the response variable. This could happen even in the case when the remaining principal components account for a large proportion of variation in the original variables . In addition, omitting low variance principal components could be problematic in the regression model, because they may be strongly associated with the dependent variable (Hadi \& Ling, 1998).

Finally, PCA raises both conceptual and statistical issues. Conceptual issues include the appropriate number of patterns derived from PCA, how patterns should be named, whether food items should be aggregated before entering the PCA, how input variables should be correctly quantified, whether or not input variables should be adjusted for total energy intake, and whether dietary patterns should be derived separately for men and women. Statistical
issues which call for a subjective judgment by the researcher include choice of the correct cut off points for eigenvalues, choice of the cut-off points for principal component loadings or component correlations in order to label a pattern, appropriate method of rotation of PCA, total variance being explained by the dietary patterns, and the interpretation of dietary patterns (Kant, 2004, Newby et al., 2004, Michels \& Schulze, 2005, McCann et al., 2001b). In addition, dietary patterns empirically derived with the use of PCA are not reproducible (Kant et al. 2004) and considered to express measures of lifestyle (Slattery et al., 1999, Maskarinec, Novotny \& Tasaki, 2000). All these issues are discussed in detail on the discussion and conclusion section of the thesis in paragraph 7.3.1.3 with the help of our systematic review of Chapter 3 and two other reviews done previously (Kant, 2004, Newby et al., 2004).

In addition, in the next paragraph we will try to investigate why PCA is different from the other miultivariate techniques and why is preferred over them in nutritional analysis.

### 2.7 A Comparison of PCA with other multivariate techniques for dietary pattern analysis

### 2.7.1 PCA and distinguishing between cases and controls

An important issue in nutritional epidemiology is the characterisation of dietary intake in relating diet to chronic disease risk. Disease risk could be estimated correctly if distinction between cases and controls is achieved according to their nutritional status. As McCann et al. (2001b) suggested, if the method used to characterise dietary intake is inaccurate with regard to measurement of the characteristics that distinguish cases from controls, diet of the cases may appear more like the diet of the controls; thus, reducing our ability to identify important dietary intake risk factors.

PCA lacks some essential features for investigating the disease structure of particular populations, as its main purpose is data reduction. Principal Component Analysis does not provide a group assessment, and would require a priori definition to distinguish between healthily and unhealthy groups of individuals according to their specific dietary intakes. But even then, PCA does not aim to obtain a clear picture of disease variation but to summarize the overall dietary intake variation of the population.

As we described in Chapter 2.4, dimensionality of a large number of dietary intake variables is reduced into two or more high-variance principal components or so-called dietary patterns. Then with the use of PCR the effect of dietary patterns on disease is estimated. As Jolliffe (2010) described in his textbook, a common problematic assumption when PCA is being employed for distinguishing between healthy and unhealthy individuals is that the correlation matrix is the same for all groups. The second problem encountered in using principal components based on a common within-group correlation matrix to distinguish between groups is that there is no guarantee that the separation between groups will be in the direction of the high-variance PCs.

For example, if the first two principal components account for a high proportion of the variance, they can also be used to provide a two-dimensional graphical representation of the data showing how good is the separation between our cases and our controls (Figures 1 and 2; pages 202 \& 203 from Jolliffe (2010)). In both figures the two groups are well separated, but in the first the separation is in the direction of the first principal component, whereas in the second the separation is orthogonal to this direction. Thus, in the second case, PCA searches for the direction showing the largest total variance and fails to distinguish between the two groups; PCA will only be useful for distinguishing between groups in the case where withinand between-group variation have the same dominant directions. If this does not occur (and in general there is no reason to expect to do so) then omitting the low-variance principal components may actually throw away most of the information concerning between-group variation (Jolliffe, 2010). Thus, 134/163=82\% of the papers (Table E) could potentially have misused PCA in this context.

Discriminant analysis is a method which is concerned with identifying well defined groups or populations of individuals. Assumptions are made about the structure of the populations, and the main objective is to construct rules for assigning individuals to populations of cases or controls according to their dietary intake. There are different forms of Discriminant analysis which could be potentially be used in nutritional epidemiological studies. Linear Discriminant Analysis (LDA) is a statistical technique similar to regression analysis except that the dependent variable is categorical rather than continuous (Jolliffe, 2010) and its aim is to predict class membership to be a case or a control based on a set of predictor dietary intake variables (McCann et al., 2001b). However, LDA is not suitable in this context because dietary intake variables are highly correlated with each other, so the assumption of independence is violated. In this case, Discriminant Analysis of Principal Components (DAPC) could provide an
alternative. In brief, DAPC relies on data transformation using PCA as a prior step to LDA, which ensures that variables submitted to LDA are perfectly uncorrelated (Jombart et AL. , 2010).

Figure 2.7.1.1. Two data sets whose direction of separation is the same as that of the first (within-group) PC.


Figure 2.7.2.2. Two data sets whose direction of separation is orthogonal to that ofthe first (within-group) PC


### 2.7.2 Other multivariate techniques

Differences in the results given by PCA with a-priori methods could be explained simply on the basis that dietary patterns derived depend solely on how well individuals are scoring on a predefined notion of diet. In this paragraph we will focus on the conceptual and methodological differences between PCA and the two other most commonly used data-driven techniques, Cluster analysis and Reduced Rank Regression. Other methods mentioned from the literature, such as cluster analysis of principal component components scores and Treelet transform are methods that do not have a wide application and they are an amalgamation of Cluster analysis and PCA, so by reviewing Cluster analysis and PCA we are partially reviewing them as well.

Main fundamental differences between PCA, Cluster Analysis and RRR rely on how dietary pattern variables are constructed and what they represent when they are associated with disease risk. In brief, PCA of data from food frequency questionnaires (FFQs) groups food items and creates a number of dietary pattern variables dependent on the degree with which their reported intakes are correlated. Principal component scores are continuous summary measures which are used as exposures in our analysis (usually the highest quintile of the score is compared with the lowest one). Cluster analysis classifies individuals into naturally existing, mutually
exclusive groups on the basis of intake of numerous foods. After clusters are defined, there is no consideration for the variation of intake within a cluster among the individuals who comprise the cluster. Clusters are used as a categorical exposure by using the largest cluster as the reference category (Reedy et al., 2010). RRR groups food intake variables and constructs dietary patterns according to the covariance matrix of specific biomarkers (taken by blood samples), nutrients, (defined by a dietary questionnaire) or other biological and epidemiological evidence which are associated with the food intake variables and are assumed to be linked with the investigated disease outcome.

So, in a simpler conceptual framework, PCA aims to explain whether there are underlying factors that explain people's diet variation; cluster analysis aims to identify different groups of individuals in the population that are consuming on average different diets; and RRR aims to find out what variation in diet is important for the development of disease with the use of an $a$ priori established biological hypothesis.

The main advantage of cluster analysis over PCA is that cluster analysis creates mutually exclusive groups which can easily be used in the analysis. On the contrary, in PCA, empirically derived patterns after a varimax rotation could be correlated with each other (see paragraph 2.4.7). Furthermore, because individuals are assigned to specific clusters according to their diets, results could be more interpretable compared to PCA.

On the other hand, PCA has the advantage of offering a technique for constructing a continuous score for individuals over a large number of foods. PCA does not only compare one group with a reference category as cluster analysis does. For example, let's assume that we derived two identically labelled dietary patterns with the use of cluster analysis and PCA; with cluster analysis we are testing if individuals who are sharing a similar "fish and vegetables" diet compared to those that sharing a "meat and potatoes" one have different disease distributions; with PCA we investigate the effect that has on disease the tendency of a population to follow a "fish and vegetables" or a "meat and potatoes" pattern. However, studies that compared PCA and cluster analysis have provided similar evidence in the nutritional literature (Bamia et al. 2007, Crozier et al, 2006, Kant et al., 2004)

The main advantage of the RRR method compared to PCA is that the pattern that is identified is associated with disease for a specific biological reason. However, when there is not a clear underlying relationship between specific markers and disease, or there are factors which are related with food intake but not with the specific markers, PCA is still claimed to be an
appropriate method to use (Kroke, 2004). Another potential advantage of RRR over PCA is that factor scores derived from RRR do not represent combinations of foods that characterize a pattern of the population diet, but combination of foods that describe a specific biomarker in the causal path with disease for that population. Finally, the RRR method as distinct from PCA and Cluster Analysis doesn't involve arbitrarily decisions on the number of factors to be derived, since this is determined by the number of biomarkers that are used as responses. Food items derived from RRR could have more plausible explanation from the foods indicated by PCA However both methods have poor repeatability of results when they used in different populations (Hoffman et al., 2004).

In conclusion, the wide use of PCA over the two other data-driven methods (Cluster Analysis and RRR) in nutritional epidemiology relies on the fact that it is simple to apply, creates a continuous numeric score and it doesn't require any plausible biological research hypothesis beforehand. PCA's long history and successful application as a data-reduction technique to different fields of research, led to the explosion of the method in the last decade in nutritional field; as indicated by our systematic review in Chapter 3 and the systematic review of Newby (2004).

In order to strengthened our arguments from our literature review in Chapter 2 and before describing our simulation study and present the results from it, we provide in Chapter 3 systematic evidence on how researchers employed PCA and PCR in years 2004-2009, and what were the key findings of the methods on the investigated health outcomes in nutritional epidemiology.

## 3 Systematic Review: Dietary patterns derived empirically from food frequency questionnaire data using PCA

3.1 Systematic Review

### 3.1.1 Aims and Objectives

3.1.2 Inclusion and Exclusion Criteria
3.1.3 Search Strategy
3.1.4 Papers Identified
3.1.5 Methods of Analysis
3.2 Summary Findings
3.2.1 Main methodological justifications for the use of PCA in dietary pattern analysis.
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"Suppose everyone had a box with something in it: we call it a "beetle." No one can look into anyone else's box, and everyone says he knows what a beetle is only by looking at his beetle. Here it would be quite possible for everyone to have something different in his box. One might even imagine such a thing constantly changing."

Ludwig Wittgenstein

### 3.1 Systematic Review

This chapter systematically review the literature on nutritional epidemiology as it relates to the use of Principal Component Analysis (PCA) for identifying dietary patterns in observational studies.

### 3.1.1 Aim and objectives

The main aim of the systematic review is to provide a presentation and discussion (see paragraph 7.3 in Discussion and Conclusion section of the thesis) of systematically collected evidence of the methodological and conceptual issues of PCA and PCR as employed in dietary pattern analysis in nutritional epidemiology. Objectives of the systematic review are:

- To identify all the relevant literature between 2004 and 2009 regarding observational studies (Tables A, B, C and D).
- To report, in each study, on how other researchers justify the use of PCA (see paragraph 3.2.1 and Table A); on the different population settings in which these studies took place and on the sample size being used (see paragraph 3.2.2 and Table B); on the type and design of questionnaires being used (see paragraph 3.2.3 and Table B); on how researchers prepared the data before entering the PCA procedure (see paragraph 3.2.4 and Table B); on the methodological decisions and numerical thresholds undertaken for deciding the number and the labelling of dietary patterns (see paragraph 3.2.5 and Table B); on the number and the total percentage of variance of the food intake variables being explained by dietary patterns (see paragraph 3.2.6 and Table B); and on the validation methods for PCA that have being used (see paragraph 3.2.8 and Table B).
- To report on the foods and food groups that were correlated highly with a dietary pattern and how these influenced the labelling of the pattern (see paragraph 3.2.7 and Table C) ; on their corresponding principal component loadings or correlations that have being used for the labelling of the "Western" and "prudent" dietary pattern (see paragraph 3.2.7 and Table D); on the level of consistency of these dietary patterns across the studies (see paragraph 3.2.7 and Tables C and D); on the major findings from studies that have explored specific effects of dietary patterns on socio-demographic characteristics (see paragraph 3.2.9 and Table C) and disease (see paragraph 3.2.10 and Table C).
- To discuss methodological issues of PCA in nutritional epidemiology (see paragraph 7.3 in Discussion and Conclusion section) and conclude on how PCA and PCR are being used in this specific context (see paragraph 7.3 in Discussion and Conclusion section).


### 3.1.2 Inclusion and exclusion criteria

The inclusion criteria that were relevant to the systematic review are defined formally below:

1. Empirical papers deriving dietary patterns with the use of PCA in nutritional epidemiology.
2. Empirical papers with the objective of illustrating, testing, criticizing or appraising PCA compared to other methods being employed for dietary pattern analysis.

The exclusion criteria for the systematic review are defined formally below:

1. Conceptual papers on methodological issues of the dietary patterns approach using PCA including think pieces and reviews of methods papers without original data (this is covered from our literature review in Chapter 2).
2. Conceptual papers comparing methods of identifying dietary patterns without original data (this is covered from our literature review in Chapter 2).
3. Studies that were using a priori dietary patters such as quality index, Mediterranean diet score or healthy eating index.
4. Studies that were using data-driven methods other than PCA.

Four further exclusion criteria were also employed:

1. Papers prior to 2004. The reason for this is that there is already a systematic review published by Newby (2004) covering previous years.
2. Articles which provided only abstracts.
3. Articles presented only to conferences.
4. Non-English language papers (there were 4 non-English language papers that were excluded).

### 3.1.3 Search strategy

The search strategy used a variety of approaches. The main part was undertaken using electronic databases, but references were also obtained by other methods, such as hand searching journals and citation searches.

Specifically, from electronic databases, to identify dietary studies that have used PCA in dietary pattern analysis, OVID was searched using the terms "(factor analysis and (diet or nutr*)) OR (principal component analysis and (diet or nutr*)) OR dietary pattern* OR ((food pattern* or eating pattern*) OR ((PCA and (diet or nutr*)). In addition, Pubmed using the terms (factor analysis OR PCA OR principal component analysis AND (diet OR nutr*)) OR ((PCA and (diet or nutr*)) was also searched.

Titles and/or abstracts from all articles retrieved from these searches were reviewed to determine whether they should be included. Reference lists from selected articles were also reviewed to locate additional papers that were not retrieved in the OVID and Pubmed search. The search was restricted to human studies reported in the English language and published up to December 2009. Data abstraction questions of the systematic review are reported in Appendix I. All the data abstracted from the papers were stored in several excel sheets by year of publication.

### 3.1.4 Papers identified

The electronic search in OVID generated 11891 references in total which, after duplicates, amounted to 9343 unique papers. Of these, 377 articles were retrieved following examination of the title and abstract, of which 163 were considered to be consistent with the inclusion and exclusion criteria and were formally included in this review (Figure 3.1.4.1). Pubmed database was further explored for papers that were not identified by the OVID search but didn't provide any additional papers.

Figure 3.1.4.1. Summary of study selection and exclusion: electronic literature searches 20042009 from OVID.


### 3.1.5 Methods of Analysis

Summary tables and graphs are used to to help explain the patterns in the data. For each data abstracted question we provide percentages on the number of papers that include information, and summary statistics (median and interquatile range). Descriptive statistics of the systematic review are presented in Table E.

### 3.2 Summary findings

Our systematic review identified 163 papers employing PCA in observational studies between 2004-2009. Figure 3.2.1 presents number of these papers by these years along with data from two other systematic reviews (41 papers between 1998-2004, and 13 between 19811997) (Kant, 2004, Newby \& Tucker, 2004). Articles from each table are summarized in their own section below. All the tables summarize why PCA is employed, how PCA is applied and how associations between diet and disease and socio-demographic population characteristics are explored through the method in each abstracted paper. Papers identified within the same cohort but reporting different research hypothesis are summarized in the same row of the table; in table A studies are organized by year of publication and alphabetical list of the author; in table B, studies are organized alphabetically by type of design and study cohort; in Table C, studies are organized alphabetically by health outcome; and in table D, studies are organized by year of publication, cohort study and alphabetical list of the author. Papers appear to be published at 2010 in our systematic review have been published electronically in year 2009, and that's why they are included. Although the term "principal component analysis" is used in this thesis, other terms, such as "factor analysis" (Agur-Collins et al., 2009), "principal components factor analysis" (Bertuccio et al., 2009) and "principal factor analysis" (Muller et al., 2009) were encountered for the same technique in papers in our systematic review. The main reason is that certain textbooks treat PCA as a special case of factor analysis, as do certain statistical computer packages (e.g. the command PROC FACTOR is SAS (SAS Institute, Inc, Cary, NC)).

Table A describes the main methodological justifications being used for the application of PCA in the abstracted dietary pattern studies; Table B summarizes the use of PCA on the populations and sample sizes being used in the articles, the type and design of questionnaires being used, how the data derived from these questionnaires are managed before entering the analysis, the criteria being used for employing PCA and articles which tried to validate the empirically derived patterns. Table C shows the number of patterns empirically derived with
the use of PCA in each article, and their associations with disease outcomes and sociodemographic characteristics. Finally, Table D presents the food elements that constitute a "prudent" and a "Western" dietary pattern and how they are labelled. For ease of reading all the tables are displayed together at the end of the thesis. A broader discussion of study findings on how PCA is commonly applied in nutritional epidemiology appears in the next ten paragraphs. Percentages and summary statistics of reasons why and how PCA is applied in the specific in nutritional epidemiology are presented in Table E at the end of the thesis.

Figure 3.2.1: Number of publications per year of empirically derived dietary patterns with the use of PCA


### 3.2.1 Main methodological justifications for the use of PCA in dietary pattern analysis

Methodological justifications for use of PCA (as described as well in paragraph 2.5) in the specific context as stated by the abstracted literature are that it is assumed that it capture additive effects of foods and nutrients on the investigated health outcome which are too small to be detected when they are examined independently ( $8.5 \%$ of the studies). PCA is also sometimes assumed to capture the effect of interaction and synergy between the different foods
and nutrients ( $15.3 \%$ of the studies). It has been claimed that PCA resolves aspects of confounding and mulitcollinearity, so allowing more accurate estimates of the effect of diet on disease ( $31.9 \%$ of the studies). It is also claimed that it can reduce the risk of chance findings arising from multiple statistical testing ( $4.2 \%$ of the studies). Public health policy is argued to be better understood in terms of dietary patterns than in terms of individual foods $(8.5 \%$ of the studies). Finally, studies have claimed more general justifications on the basis that PCA could explain the complexity and provide a better assessment of the overall diet ( $18.4 \%$ of the studies) (Table A and E).

### 3.2.2 Populations and sample size being used

Principal Component Analysis was employed in a variety of observational studies for diverse populations and settings, in every continent, including several countries in Europe, North America including Canada, Latin America (Argentina, Brazil, Mexico, Costa Rica, Uruguay), Middle East (Iran) and Asia (China, Japan, Singapore, Korea, Bangladesh), Caribbean (Jamaica, Puerto Rico), Australia, Canada and Africa (Botswana, Mauritius). Sample size of the populations being used varied from 115 (Custodio das Dores et al., 2007) to 492382 (Flood et al., 2008) (Table B).

### 3.2.3 Details on type and design of questionnaire being used

Quantitative food-frequency questionnaires (FFQs) and semi quantitative food frequency questionnaires (SFFQs) (designed for measuring food consumption quantitatively and qualitatively) were employed as the primary assessment method in $93.2 \%$ of the studies. Some studies used diet history questionnaire (1.8\%) (EPIC, DHQ) (Masala et al., 2007, Cottet et al., 2009, Robinson et al., 2009, Yannakoulia et al., 2008, Okubo et al., 2008, Waijers et al., 2006) , 24 hour recall (2.4\%) (Cui et al., 2007, Kesse-Guyot et al., 2009, Hamer \& Mishra, 2010, Kim et al., 2007), 48 hour recall (Mikkila et al., 2005) and dietary records (1.8\%) (Balder, Goldbohm \& van den Brandt, 2005, Cuco et al., 2006, McNaughton et al., 2007, Perrin et al., 2005, Newby, Muller \& Tucker, 2004b) to assess diet. Individual food and nutrient intakes were derived from these methods and most studies collapsed the original measured dietary items into a smaller number of input variables, usually food groups, for entry into the principal component analysis ( $28.4 \%$ of the studies). Food items in the questionnaire varied from 11 (Takaoka \& Norback, 2008) to 255 (Bakolis et al., 2010, Pala et al., 2006) (median value: 92, IQR: 21-204) and food groups derived varied from 13 (Romaguera et al., 2008) to 74 (Ambrosini et al., 2008) (median value: 38, IQR: 19-69). Scale of the FFQ being used varied
from 5 (Wiles et al., 2009, Sadakane et al., 2008, Keskitalo et al., 2008, Shimazu et al., 2007) to 10 (Bakolis et al., 2010, Flood et al., 2008, Campbell, Sloan \& Kreiger, 2008, Yang, Kerver \& Song, 2005) point scales (median value: 7; IQR:5-10). Four studies used macronutrient and/or micronutrient intakes in the analysis rather than foods or food groups (Bertuccio et al., 2009, Edefonti et al., 2008, De Stefani et al., 2008a, Corrao et al., 2004) (Table B and E). Finally, $64 \%$ of the studies were cross-sectional, $19 \%$ were cohort, $17.1 \%$ were case-control studies and $2.4 \%$ were clinical trials (Table E).

### 3.2.4 Preparation of data before the application of PCA

Food intake variables may be measured in frequency (servings), weight (grams), or daily percent energy contribution. Conversions to grams/d or grams/week and standardization of the food intake variables was reported in $16 \%$ of the studies. Moreover, in 10 studies food intake variables were adjusted for energy intake by the residual method before entering the PCA procedure (Willett, 1998) (Table B and E).

### 3.2.5 Empirical derivation and labelling of the dietary patterns

Decisions for retaining the number of principal components were based on the following criteria. Cut - off points for eigenvalues to decide the number of principal components ranged from 1 (Kaiser Criterion) to 3 (Kim et al., 2008) (median value: 1.6, IQR: 1-2). In addition, $49.1 \%$ of papers used a scree plot and $42.9 \%$ of the studies derived patterns according to their principal component interpretability. In one paper Van de Voet's test (DiBello et al., 2008) was also used for deciding the number of factors.

Decisions for retaining the label of principal components were based on the following criteria; Cut off points for component loadings or component correlation coefficients for deciding which foods constituted the dietary pattern ranged from 0.15 (Hughes et al., 2009) to 0.6 (Park et al., 2005) (median value: 0.3, IQR: 0.3-0.4). When principal component analysis was employed on a number of nutrients (3 studies), cut off points for component loadings ranged from 0.39 (De Stefani et al., 2007) to 0.63 (Edefonti et al., 2008). Principal components were rotated with the use of varimax (orthogonal) rotation for better interpretability of the patterns in $52 \%$ of the studies. Only four studies used promax (oblique) rotation (Kim et al., 2008, He et al., 2009, Lau et al., 2008, De Stefani et al., 2007) (Table B and E).

### 3.2.6 Number of dietary patterns and percentage of total variance being explained by them

The number of empirically derived dietary patterns ranged from 2 to 10 (median: 3, IQR: 24) and the total percentage of variance being explained by the dietary patterns ranged from 11.2\% (Hooper et al., 2010) to 88\% (Romaguera et al., 2008) (median: 24\%, IQR: 19.9-31.3). Not all the papers provided information of the total variance being explained by the dietary patterns.

### 3.2.7 Foods and food groups that were deemed to constitute a dietary pattern and how

 these influenced the labelling of the patternPrincipal Components were labelled quantitatively according to:

- Food items with the highest principal component loading such as "coffee", "bread".
- Food groups with the highest principal component loading such as "vegetable", "sweets", "meats", "alcohol", "fruit", "bread eaters", "less fish", "confectionary", "plant-based" (3 studies), "animal food", "salad vegetables", "vegetable-soy", "meat-sweet", "stew".
- Combination of food items and food groups such as "vegetables, fruit and milk", "fruit, salad, cereals, and fish", "meat and fast food", "fruit and milk", "fish and sauce", "potato and fish", "cereals and legumes", "alcohol and butter", "fruits and vegetables", "whole food", "potato and fish", "cakes and sweets", "fruits", "coffee and dairy", "meat-starch", "sweets and soft drinks", "British meat and two vegetables", "fruit-rich", "meat-rich", "refined-grains", "meet dim-sum", "sweetened beverages and sugars", "pasta and meat", "olive oil and salad", "sweet and dairy", "pork processed meat and potatoes", " cooked vegetables" and "green".
- Descriptions of dietary composition of the food items that were highly correlated to the pattern such as "carbohydrate", "antioxidants", "low-fat/low-sugar", "high-fat", "high-protein/high-fat", "phytoestrogen-rich", "animal proteins", "vitamins and fibre" and " unsaturated fats".
- Descriptions of the way that foods were cooked such as "processed" (11 studies) with high positive principal component loadings or correlation coefficents between the
pattern and processed meat products, white bread, French fries, salty snacks, and sugar-sweetened drinks and high negative principal component loadings or correlation coefficients between the pattern and oily fish, high-fibre breakfast cereals, and lean fish.

Furthermore, dietary patterns were labelled qualitatively according to foods or food groups which were assumed to provide a degree of benefit. The majority of the studies decided to label their empirically derived dietary patterns in this way, as "prudent" (43 studies). The foods comprising a "prudent" diet varied from one study to another but the following characteristics appeared as part of a "prudent" diet in more than two studies: high intakes of vegetables, cheese, olive oil, fruit, wholemeal bread, non-fried fish, poultry, rice/pasta, beans, eggs, seafood, yoghurt and breakfast cereals and low intakes of white bread, roast potatoes/chips, red/processed meat, full-fat milk, full-fat spread, crisps, confectionery, sugar, tea/coffee and Yorkshire puddings/ pancakes, tinned vegetables, cakes and biscuits and soft drinks (Table D). Other dietary pattern names which were highly positively or negatively correlated with the same food items as the "prudent" pattern were entitled "healthy" (30 studies), "health conscious" ( 5 studies), "health aware" and "heart healthy". Under the same framework, patterns were labelled according to how unhealthy they were assumed to be, such as "junk food" and "junk" (3 studies), "fast food" (2 studies) "avoidance" and "unhealthy" (Table C).

Another way of labelling a pattern was according to the degree which was positively correlated with foods or food groups which were linked with a specific lifestyle. Specifically, patterns were labelled "Western", "Western-like", "Western-type", "Macho" in 65 studies when they were strongly positevely correlated with cured and red meats, white bread and rolls, chocolate, margarine, butter desserts, potatoes, sweets, pizza, soft drinks, French fries, coffee, alcohol, high-fat dairy products, hamburger, eggs, bacon, mayonnaise, doughnuts, and negatively correlated with rice, vegetables and low fat milk (Table D). In addition patterns were labelled "vegetarian" ( 6 studies) if they included negative loadings for food items that were avoided by vegetarians, principally meats; "diet", if they included foods items that were consumed mainly by people who were on a diet; "sweet tooth" and "drinker" (6 studies) (Table C).

Other qualitative labels referred to cultural or geographic descriptions of dietary intake, such as "traditional" (39 studies), "modern", "Iranian", "Andean-like", "Mediterranean",

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"ethnic", "yellow earth", "green water", "adopter", "new affluence", "oriental", "native
Mexican", "urban/rural", "seasonal", "Korean", "Swedish", "healthy American",
"Japanese", "traditional southern", "traditional Finnish", "traditional Dutch" and
"canteen".
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Finally, some articles chose to name the principal components as "Component 1 and 2" and "Factor 1 and 2" (Table C).
(Note: all counts refer to separate study populations and not to separate reports from the same study population. We counted only patterns that were observed in more than one population).

### 3.2.8 Validation methods being used after PCA for confirmation of the derived patterns

Very few validation studies of eating patterns have been performed. Five of the validation studies considered the stability of patterns over time, an indicator of reproducibility, by comparing factor solutions from an FFQ at different time points (Northstone, Emmett \& Rogers, 2008, Lau et al., 2008, Fung et al., 2005, Khani et al., 2004, Togo et al., 2004). Confirmatory factor and maximum likelihood analysis for principal components were used to test the internal validity of a pattern in 8 of the studies. Internal validation of the empirically derived principal components was also performed by randomly splitting the study sample and repeating the principal component analysis to the two sub-samples in 8 of the studies; by deriving patterns separately for men and women and comparing the results in 3 of the studies; by using Cronbach's alpha reliability coefficient which evaluates to what extent food items measure the same underlying content when the food items are combined into a scale in 5 of the studies; by using Barlet's test of sphericity (1 study) which tests the null hypothesis that our sample was randomly drawn from a population in which the correlation matrix was an identity matrix (a matrix full of zeros, except for ones on the main diagonal); by employing the Kaiser -Meyen-Ollkin test ( 2 studies) which tests whether the partial correlation coefficients among our food intake variables are small; by constructing simplified dietary patterns according to Schulze method (1 study) (see paragraph 4.9); and by employing a different method of rotation (2 studies) (Table B).

### 3.2.9 Associations between Principal Components and socio-demographic characteristics

Table C presents results from 60 studies ( $36.8 \%$ of the overall number fo studies) that have examined associations between empirically derived dietary patterns with the use of PCA and
socio-demographic characteristics. Studies reported a high "fruits, vegetables and oily fish" pattern consumption ( labelled as "prudent" or "health conscious" or "vegetable or "fruit and vegetable" or " Mediterranean" or "healthy" or "low-fat" or "vegetable-soy" or " fruit, salad, cereals, and fish" and were positively associated with women, older age, vitamin and supplement use, height, income, social class, strenuous exercise, education, urban living, partnership status, non-smoking status, non-obese status. A high "red meat, processed foods, sugary items and soft drinks" pattern of consumption (labelled as "Western" or "processed" or "processed meat" or "organ meat and fast food" or "meat and French fries" or "alcohol/meat" or "unhealthy" or "high energy" or "meat-dim sum" or "highmeat" or "fat-reduced and diet-foods" or "fat and sugar" or "junk" or "vegetables and fruit and milk") was inversely associated with age, exercise, education, social class, height and positively associated with being male, BMI, smoking and drinking, height urban and rural residence. A "convenience foods" pattern consumption was inversely related to age and positively related to higher education in both genders and a "starch, sauces, and vegetables" pattern consumption was associated with high education and an urban residence (Kesse-Guyot et al., 2009). People consuming a "carbohydrate" pattern showed lower use of vitamins and were less likely to be overweight or obese. An "alcohol" pattern was inversely associated with weight gain (Jackson et al., 2009). Men in the highest tertile of the "sugary foods and sweet baked goods" pattern were more likely to be current smokers and were less likely to have tertiary education (Jackson et al., 2009, De Stefani et al., 2009). A "traditional" pattern consumption based on a UK population was more common in men, in couples and was associated with higher alcohol use (Robinson et al., 2009). People consuming a "traditional / Iranian" dietary pattern score were older, slightly more physically active (Esmaillzadeh \& Azadbakht, 2008b). People scoring high in a "Japanese" pattern were less likely to be smokers and drinkers (Hirose et al., 2007). Subjects with a high score for the "drinker" pattern were younger and were more frequently current smokers. "Stew" pattern was inversely correlated with education (Mizoue et al., 2006). Higher scores for the "traditional Dutch dinner" dietary pattern were associated with women who had a lower level of education, were current smokers, and were more overweight (Waijers et al., 2006). The "traditional Korean" (Yang, Kerver \& Song, 2005) dietary pattern was negatively associated with length of residence in the U.S. for both men and women. Older women and with no children had high scores for the "canteen" pattern (Sieri et al., 2004) (Table C).

### 3.2.10 Major findings between Principal Components and health outcomes

Table C presents results from 120 studies ( $73.6 \%$ of the total number of studies) that have examined potential associations between empirically derived dietary patterns with the use of PCA and disease outcome. Dietary patterns to major health outcomes which are examined will be presented briefly as evidence of the potential predictive consistency of the PCA method only.

Derived patterns were examined in relation to many different outcomes, including indicators of cardiovascular or coronary heart disease; different type of cancers (gastric, breast, prostate, colorectal, bladder, gastric, lung, pancreatic, laryngeal, endometrial); overweight and obesity; metabolic syndrome and its components, including hypertension, blood pressure, cholesterol, blood glucose, and blood insulin measures; type 2 diabetes; all-cause mortality; depression; respiratory conditions including asthma, chronic bronchitis, wheeze, cough with phlegm and COPD; Preeclampsia; VTE (venous thromboembolism); Behavioural scores; spina bifida; myocardial infarction; Barrett's oesophagus; school attainment; dyslipidemia; Crohn's Disease, constipation; overall health status; and stroke.

## Overall mortality

An increase in the pattern score which measures the adherence to the "plant-based", (Bamia et al., 2007), "fruit-rich" (Cai et al., 2007) and a "healthy traditional" (Waijers et al., 2006) diet was associated with a lower overall mortality. An association of the "pasta $\&$ meat" (Bamia et al., 2007), "meat-rich" (Cai et al., 2007) pattern with increased overall mortality was suggested by the literature. Furthermore, "prudent" (Masala et al., 2007), "Mediterranean" (Waijers et al., 2006), "meat, potatoes, legumes and bread" and "vegetables, fruits and dairy products" (Hoffmann et al., 2005) patterns were not associated with overall mortality.

## Cardiovascular Disease (CVD)

A "whole grains and fruit" (Nettleton et al., 2009) and a "prudent" pattern were associated with a lower risk of cardiovascular disease (Heidemann et al., 2008). On the other hand, a "Western" (Heidemann et al., 2008), "animal food" (Shimazu et al., 2007) and "healthy" (Esmaillzadeh \& Azadbakht, 2008a) patterns were associated with a greater risk of incident CVD. In addition, a "component 1" pattern which loaded mainly on low-fat products like fish,
vegetables, legumes, greens, and salads, as well as olive oil was associated with lower likelihood of having increased burden of CVD (( et al., 2007a).

## Myocardial Infarction (MI)

A "vegetable" (DiBello et al., 2008), and "prudent" (Iqbal et al., 2008) patterns were associated with a significantly decreased risk of myocardial infarction. .A "western" pattern showed a U-shaped association with Acute Myocardial Infarction (Iqbal et al., 2008). "Healthy" and "alcohol" patterns were positively associated with increased risk of MI (Akesson et al., 2007).

## Type 2 Diabetes mellitus including Type II diabetes

Dietary patterns characterized by high intake of "whole grains, fruit, nuts/seeds", "green leafy vegetables", "low-fat dairy" (Nettleton et al., 2008b) and "prudent" (Montonen et al., 2005) characteristics of diet were associated with lower diabetes risk. Food patterns characterized by high intake in "high fat and sugar" (McNaughton, Mishra \& Brunner, 2008) and "tomatoes, beans, refined grains, high-fat dairy, and red meat" (Nettleton et al., 2008) were associated with increased risk of type 2 diabetes. "Western" and "prudent" (Imamura et al., 2009, Fung et al., 2004a) pattern scores were not associated with Type 2 diabetes. Finally, a "Westernized breakfast" was inversely and a "seafood" dietary pattern was positively associated with A1C concentrations related to type 2 diabetes (Nanri et al., 2008b).

## Metabolic Syndrome

Subjects in the highest quintile of a "healthy" dietary pattern score had a lower odds ratio for the metabolic syndrome. Consumption of a "Western" (Esmaillzadeh \& Azadbakht, 2008a) and "sweets" (Noel et al., 2009) dietary pattern was adversely associated with incident metabolic syndrome. "Korean traditional", "Western" (Kim et al., 2007) and "prudent" (Lutsey, Steffen \& Stevens, 2008) patterns were not statistically significant associated with metabolic syndrome.

## Types of Cancer

## Breast

A "prudent" diet was inversely (Hirose et al., 2007, Agurs-Collins et al., 2009) not (Fung et al., 2005, Kroenke et al., 2005, Robinson et al., 2004), and positively associated (Murtaugh et al., 2008) with breast cancer. A "Western" pattern was positively (Cottet et al., 2009, Murtaugh et al., 2008, Wu et al., 2009, Ronco et al., 2006) and not associated (Fung et al., 2005, Agurs-Collins et al., 2009, Kroenke et al., 2005, Robinson et al., 2004) with breast cancer. "Traditional", "healthy", "stew" (Ronco et al., 2006), "vegetables" (Takata et al., 2007), "salad vegetables" (Sant et al., 2007), "salad-sauce-pasta/grain" (Tseng et al., 2008), "vegetables" (Takata et al., 2007) diets were significantly protective against breast cancer.

## Colon

An increased risk of colon cancer and distal adenoma was suggested with higher "Western" (Wu et al., 2004a, Meyerhardt et al., 2007) and "pork, processed meats and potatoes" (Dixon et al., 2004) pattern scores. Higher "prudent" pattern scores were only weakly and non-significantly associated with decreased risk of colon cancer or distal colon adenoma ( Wu et al., 2004a).

## Colorectal

A "Mediterranean" pattern significantly reduced colorectal cancer (Cottet et al., 2005). High scores of "meat-eaters" (Kesse, Clavel-Chapelon \& Boutron-Ruault, 2006), "pork, processed meats and potatoes" (Dixon et al., 2004), "meat and potatoes" (Reedy et al., 2010) and "red meat" (Flood et al., 2008) patterns had higher risk of developing colorectal cancer risk. "Healthy", "fruits and vegetables", "drinker", "snacks" (Kesse, ClavelChapelon \& Boutron-Ruault, 2006, Cottet et al., 2005), "meat-dim sum", "vegetable-fruitsoy" (Butler et al., 2008), "healthy", "Japanese" and "animal food" (Mizoue et al., 2005) patterns were not associated with colorectal cancer risk.

## Gastric

"Healthy", "prudent" (Kim et al., 2004, De Stefani et al., 2004) "fruit, salads, vegetables, dairy products, fish and meat" (Bastos et al., 2010) "vitamins and fiber", "vegetable", "mixed" (De Stefani et al., 2004) patterns were associated with decreased risk of gastric cancer. A positive association between gastric cancer risk and the "animal products" ,"starch-rich" (Bertuccio et al., 2009) ,"Western" (Campbell, Sloan \& Kreiger, 2008) and
"starchy" (De Stefani et al., 2004) dietary pattern was observed. A "Western" (Kim et al., 2004) dietary pattern was not associated with risk of gastric cancer.

## Lung

"Antioxidants" (De Stefani et al., 2008a) 'salad vegetables" and 'sweet foods"' (Balder, Goldbohm \& van den Brandt, 2005) pattern were all inversely associated with risk of lung cancer. A "high-meat" pattern was associated with a strong increase in risk of lung cancer (Dixon et al., 2004). The "carbohydrates" pattern was not associated with risk of lung cancer (De Stefani et al., 2008a).

## Prostate

An increased risk for prostate cancer was observed with a higher intake of a "Western" (Ambrosini et al., 2008, Arkkola et al., 2008) and a "Southern" pattern (Tseng et al., 2004). There were no associations between "meat \& potatoes" (Muller et al., 2009) "red meatstarch" (Tseng et al., 2004) "Western" (Wu et al., 2006), "vegetable", "health-conscious" (Ambrosini et al., 2008), "Mediterranean", "fruit \& salad" (Muller et al., 2009), "prudent" (Wu et al., 2006), "vegetable-fruit" (Tseng et al., 2004) and "healthy and carbohydrate" (Jackson et al., 2009) patterns with overall prostate cancer risk.

## Other

The "prudent" pattern was directly associated with risk of bladder cancer (De Stefani et al., 2008a). "Pattern 5" (drinker) and "pattern 6" (western) were directly associated with risk of laryngeal cancer whereas the "pattern 2" (healthy) was protective. (De Stefani et al., 2007) . No associations were observed between the "prudent" and "Western" pattern and the risk of pancreatic cancer (Michaud et al., 2005). Intakes of "fruits and vegetables" pattern were associated with a reduction in risk of pancreatic cancer (Nkondjock et al., 2005). A "healthy" pattern was not significantly associated with decreased risk of Renal Cell Carcinoma (all cancer and kidney cases) (Rashidkhani et al., 2005). A "plant-based" diet had higher ovarian cancer risk (Chang et al., 2008).

## Respiratory and allergic symptoms

A "vegetarian" dietary pattern was positively associated with asthma (Bakolis et al., 2010). There were no evidence that "fish, fruits and vegetables" (Hooper et al., 2010b), "health
conscious" (Shaheen et al., 2009) "western" and "prudent" (Bakolis et al., 2010) patterns were associated with asthma. A "Western" pattern was associated with an increased risk of reporting frequent asthma attacks and a "nuts and wine" with decreased risk (Varraso et al., 2009). A "prudent" pattern was positively associated with chronic bronchitis (Bakolis et al., 2010). An "urban" component of diet was strongly associated with positive skin tests after adjusting for place of residence (Hooper et al., 2008). The "prudent" pattern was inversely associated and a "Western" pattern was positively associated with the risk of newly diagnosed COPD (Varraso et al., 2007a, Varraso et al., 2007b). The "meat-dim sum" pattern was positively associated with new-onset cough with phlegm (Butler et al., 2008). A pattern including "fast food, juice and soft drinks" was related to wheeze and respiratory infections (Takaoka \& Norback, 2008). Univariately, a "health conscious" pattern was positively associated with eczema and a "processed" pattern was negatively associated with atopy (Shaheen et al. 2009).

## Obesity

"Healthy" (Okubo et al., 2008, Newby et al., 2004a), "prudent" (Murtaugh et al., 2007, Paradis et al., 2009, Newby et al., 2006) patterns were significantly associated with a lower risk of obesity. Individuals with a high consumption of "Western" (Murtaugh et al., 2007, Paradis et al., 2009), "Japanese traditional" (Okubo et al., 2008), "pasta \& meat" (Pala et al., 2006), "animal foods" (Shin, Oh \& Park, 2007) and "fish and sauce" (Craig et al., 2010) pattern were more likely to be obese. No significant associations between "green", "sweet" and "traditional" patterns and obesity were found (Togo et al., 2004).

## High-density lipoprotein (HDL) Cholesterol

Serum HDL cholesterol was inversely associated with "health aware" (Hamer \& Mishra, 2010), "sweets" (Newby et al., 2004b) and "Western" (Deshmukh-Taskar et al., 2009) dietary pattern. "Meat", "Western", "vegetable" and "protein and alcohol" (Hamer \& Mishra, 2010, Sadakane et al., 2008, Newby et al., 2004b) patterns were associated with higher total HDL cholesterol.

## Mental and Behavioural Health

Improved behavioral scores were significantly associated with a "healthy" pattern (Oddy et al., 2009). Participants in the highest tertile of the "whole food" pattern had lower odds of
depressive symptoms than those in the lowest tertile. Patterns labeled as "sweets" and "meat and products" were positively associated with anxiety score in females (Yannakoulia et al., 2008). Child Behaviour Checklist for mental health scores was significantly associated with the "Western" dietary pattern. A "junk" food pattern was negatively associated with school attainment (Feinstein et al., 2008). In addition, high consumption of "processed food" was associated with increased odds of depressive symptoms (Akbaraly et al., 2009). There was little evidence to support an association between "junk food" intake and overall behavioral difficulties or other sub-scales of the childhood behavioral problems (Wiles et al., 2009).

## Other health outcomes

A "balanced" pattern was associated with decreased risk and an "animal protein" with increased risk of hypertension (Chen et al., 2006). Strong adherence to the "health-conscious" dietary pattern was inversely associated with Barrett's esophagus (Kubo et al., 2008). A "traditional Western" in girls was positively associated and a "prudent" was inversely associated with Chron's Disease (D'Souza et al., 2008). There were no consistent association of "vegetable-fruit", "potato-sweats and meat" and "alcohol-snacks" patterns with Actinic Keratoses acquisition (Hughes et al., 2009). A pattern characterized by high consumption of "fish and olive oil" and low intake of "red meat" was positively associated with lumbar spine bone mineral density (Kontogianni et al., 2009). Individuals who had a high consumption of a "health conscious" diet had lower sperm DNA damage. Furthermore, sperm concentrations were much higher in men who strongly adhere to the "traditional Dutch" dietary pattern (Vujkovic et al., 2009a). A significantly increased risk of spina bifida was observed for offspring in mothers with a weak use of the "Mediterranean" dietary pattern (Vujkovic et al., 2009b). A dietary pattern characterized by high consumption of "whole grains, fruit, vegetables", and "low-fat dairy foods" was associated with lower spot urine collection (Nettleton et al., 2008b). The "prudent" pattern was inversely and the "Western" pattern was positively associated with plasma concentrations of CRP (CRP are markers of endothelial dysfunction) (Lopez-Garcia et al., 2004). A "Vegetables, plant foods and vegetable oils" pattern associated with lower risk and a "processed meat, salty snacks, and sweet drinks" pattern with increased risk of Pre-eclampsia (Brantsaeter et al., 2009). A "high-dairy, highfruit and vegetable, high-starch, low-alcohol" pattern was significantly and inversely associated with glucose tolerance abnormality. Additionally, a "dairy products and fruits
and vegetables" pattern was associated with decreased risk of developing a glucose tolerance abnormality ( $\mathrm{P}<0.05$ ) (Mizoue et al., 2006).

A detailed discussion and overall conclusions of these findings are given at paragraph 7.3. Moreover, after being informed with the help of our literature (Chapter 2) and systematic (Chapter 3) review on the application of PCA procedure in nutritional epidemiology we proceed to Chapter 4, where PCA is critically evaluated and compared with an Exhaustive Single food Analysis with the use of a Monte Carlo simulation study.

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"The truth content of our theories, even the best of them cannot be verified by scientific testing, but can only be falsified."

Karl Popper

### 4.1 Overview

Our aim of this chapter is to describe the methodology behind our simulation study. A simulation study as Maldonado and Greenland (1997) suggested is a study which repeatedly answers a "What if?" question. In our case, this "What if?" question asks about the performance of two statistical methods, an exhaustive single food analysis (ESFA) and a Principal Components Regression (PCR).

Our main premise is that in order PCR to characterize the diet associated disease risk better than anyone food, then PCR should identify combinations of foods in the population which are causally linked with disease. We set out to investigate whether PCR really performs better in these respects than an analysis of individual food intakes.

In order to test this hypothesis, we created a hypothetical population in which food intake derived from a food frequency questionnaire could be related to disease under the conditions that we controlled for and in a way that we could specify in terms of relative risks. In our simulation model, a number of randomly selected foods that are not highly correlated with each other and a "Western" dietary pattern (a selected number of foods that are correlated with each other) are assumed to be in the causal path with disease.

Our objective is to investigate whether analysing each individual food on the FFQ separately, a process we refer to as an exhaustive single food analysis (ESFA) is as good as a method as constructing posterior dietary patterns and effectively identifying combination of foods that are causally linked with disease in this hypothetical population with the use of PCR.

### 4.2 Simulation study

### 4.2.1 Principal of Bootstrapping

In literature, the term bootstrap is attributed to Rudolph Erich Raspe and its story on the adventures of Baron Munchausen. The Baron in order to prevent himself from drowning in a lake pulled himself out of his bootstraps. This term is used when someone wants to describe the impossible. In statistics, bootstrap is a special case of the Monte Carlo simulation method, where re-sampling techniques are used to assign measures of accuracy to statistical estimates and was originated by Efron in 1979 and developed by Efron and Tibshirani (1993), as a way to tackle inaccurate and complex mathematical procedures. We will give a detailed description of the method as employed in our simulation study.

### 4.2.2 Simulating diet

Individuals who took part in the F.L.A.G survey (Data-set 1 -see paragraph 4.8.1) and ECHRS UK study (Data-set 2 -see paragraph 4.8.2) are two representative random samples of the populations living in Greenwich and Ipswich/Norwich. In each study, we treat the sample we have as if it were itself a reference population. We sample from the sample we have with replacement. Sampling with replacement means that some of the individuals will be part of our new sample more than once and some others will not be part at all. This imitates the way we would have drawn a new sample if we wanted to conduct another study for the same reference population. These samples of our sample are called bootstrap samples.

More formally, because PCA, as applied in nutritional epidemiology, depends on the correlation matrix of all dietary intake variables in particular we wanted to ensure realistic correlation structure in ours bootstrap simulations. For this reason we used real data sets. The real data-set that we used is referred to here as the reference data-set. For each individual, dietary consumption was recorded with the use of a food frequency questionnaire (FFQ) over a number of foods and beverages and dietary intake variables were constructed for our two reference data-sets. Food intake variables in the dataset were standardised to have a mean 0 and a standard deviation of 1 ; whereby a negative value corresponds to a smaller-than-average value of dietary intake; a zero value corresponds to an average value; and a positive value corresponds to a larger-than-average value. To simulate a dietary data-set (corresponding to the bootstrap sample of individuals as described above) with a realistic correlation structure between food intake variables we sampled with replacement from the reference data-set. When we sample with replacement, sample values of food intake are independent and it is like sampling from an infinite population.

### 4.2.3 Simulating disease from randomly selected food items and a simplified dietary pattern

We then simulated the presence or absence of disease assuming a logistic model for the disease risk. This model has been widely used in previous simulation studies in epidemiology (Fewell et al., 2007, Pastor \& Guallar, 2001, Peduzzi et al., 1996). In each simulation we assumed that disease risk depended on a linear combination of $m$ food intakes from the FFQ. Suppose these foods are indexed $i_{1}, i_{2}, \ldots, i_{m}$, and absolute food intakes $x_{i_{1}}, x_{i_{2}}, \ldots, x_{i_{m}}$ are standardised to have zero mean and unit standard deviation. We assumed a logistic model for the disease risk, $p$ that is

$$
\begin{equation*}
\log \frac{p}{1-p}=a+b_{1}+b_{2} x_{2}+\ldots+b_{m} x_{m} \tag{4.2.3.1}
\end{equation*}
$$

or solving (4.2.3.1) for $p$, we have that

$$
\begin{equation*}
p=\frac{1}{1+e^{-\left(a+b_{1} x_{1}+b_{2} x_{2}+\ldots . b_{m} x_{m}\right)}} \tag{4.2.3.2}
\end{equation*}
$$

We chose the constant $a$ so that the baseline risk at the average intake of all foods was 0.15 , i.e. $a=\ln (0.15 /(1-0.15))$. Constants $b_{1}, b_{2}, \ldots, b_{m}$ were chosen so that the odds ratio per standard deviation of food intake was 1.5 or $1 / 1.5$ depending on whether the food was assumed to increase or decrease the risk of disease, i.e. $b_{j}= \pm \ln (1.5)$. Combing data from all the dietary pattern studies in our systematic review and from Newby (2004) we observed that the median value of statistical significant odds ratio was 1.38 (IQR: 0.77-1.99) (see table C at the end of the thesis and Newby (2004)). We decided to use a slighter higher odds ratio, in order to increase the power of our simulation study.

The value of $m$ was chosen to take two different values for each of the two data-sets; around one in seven of the total number of foods on the two FFQs, that is $m=30$ (of 217) in Data-set 1 , and $m=10$ (of 74) in Data-set 2; and one in twenty of the total number of foods on the two FFQs, that is $m=10$ (of 217) in Data-set 1 , and $m=4$ (of 74) in Data-set 2. The $m$ foods were chosen in two different ways.

First, they were randomly chosen in each simulation. We considered three models of this kind: in Model 1, all $m$ foods were assumed to be protective; in Model 2, $m / 2$ foods were assumed to increase the risk of disease, and the other $m / 2$ were assumed to have protective effects; in Model 3, all $m$ foods were assumed to increase the risk of disease. Results of these simulations tell us about the average performance of different methods when we do not restrict a priori the combinations of foods that might be important for disease risk. Furthermore, in our simulations, we don't assume a priori that foods must be highly correlated with each other in order to be associated with disease risk.

Second, they were predetermined in each simulation to be foods making up a simplified "Western" dietary pattern (see paragraph 4.9), which was assumed to be positively associated (Model 4) and negatively associated (Model 5) with disease risk. These foods are listed in Table 4.9.2.1, and were chosen as the food intakes with the highest positive component correlations on a "Western" dietary pattern obtained using PCA from the original, reference
data-sets ( 30 foods for Data-set 1 paragraph 4.9.2, 10 foods for Data-set 2-see paragraph 4.9.3). Results from models 4 and 5 might be expected to favour PCA as a means of identifying dietary associations, since the model is based on a principal component in the populations under investigation.

### 4.2.4 Bernoulli trial

Hence, having simulated the food intake data for a new individual for the sample, we then calculated the probability of the outcome, $p$, for each individual using equation 4.2.3.2. We determined whether or not the individual has the disease by generating a uniform random number between 0 and 1 , and observing whether it was less than $p$. If the uniform generated number was less than $p$ then the individual had the disease, but otherwise not. In this way, a Bernoulli trial for each individual in the data-set was randomly simulated and a hypothetical binary variable (our hypothetical disease) was generated for our data.

### 4.3 Principal Components Regression (PCR) and Exhaustive Single Food Analysis (ESFA) of the simulated diet-disease dataset

### 4.3.1 Description of PCR and ESFA procedures

The simulated dietary data were subjected to a PCA with a fixed number of principal components, with the use of the correlation matrix of the bootstrap sample of the reference dataset (see Chapter 2), according to the way that was described in the nutritional literature (see Chapter 3); majority of studies that employed PCA derived 2 to 10 principal components (median value was 3 , IQR: 2-5) which were varimax rotated. A principal component score was calculated for each varimax rotated dietary pattern for each individual in the sample to represent the individual's level of intake for the pattern. Resulting rotated principal component scores (dietary patterns intake) were investigated for their associations with disease using a logistic regression adjusted for total energy intake. Hence, let's assume that $Y_{i}$ denote the disease outcome, (total.energy) ${ }_{i}$ the total energy intake, and $Y_{i j}^{*}$ the $j^{\text {th }}$ pattern for the $i^{\text {th }}$ individual of our sample of $n$ individuals, then
$\log \left[\frac{E\left(Y_{i} / Y_{i j}^{*}\right)}{1-E\left(Y_{i} / Y_{i j}^{*}\right)}\right]=\beta_{o j}+\beta_{1 j} Y_{i j}^{*}+\beta_{2 j}(\text { total.energy })_{i}, i=1, \ldots, n ; j=1, \ldots, p$
where $\beta_{0 j}, \beta_{1 j}, \beta_{2 j}$ are the regression coefficients of the $j^{\text {th }}$ pattern intake variable.

For comparison, we considered the results of analysing each individual food on the FFQ separately in relation to disease risk, a process we refer to as an exhaustive single food analysis (ESFA), adjusted for total energy intake. This is a univariate or independent screening approach as we know from, for example, SNP screening in statistical genetics (Laird, 2011; page 111). Hence, let's assume that $Y_{i}$ denote the disease outcome, (total.energy) ${ }_{\mathrm{i}}$ the total energy intake and $X_{i j}$ the $j^{\text {th }}$ food intake variable for the $i^{\text {th }}$ individual of our sample of n individuals, then

$$
\begin{equation*}
\log \left[\frac{E\left(Y_{i} / X_{i j}\right)}{1-E\left(Y_{i} / X_{i j}\right)}\right]=\beta_{o j}+\beta_{1 j} X_{i j}+\beta_{2 j}(\text { totalenergy })_{i} i=1 \ldots n, j=1 \ldots k \tag{4.3.1.2}
\end{equation*}
$$

where $\beta_{0 j}, \beta_{1 j}, \beta_{2 j}$ are the regression coefficients of the $j^{\text {th }}$ food intake variable.

Observational nutritional studies are strongly advised to be adjusted for energy intake (Willet, 1998, Jakes et al, 2004). As pointed out by Willet, adjustment for total energy intake should be considered because the level of intake might be a risk factor, might distort the effect of a food or a nutrient on the potential outcome, and the variation of nutrient intake between individuals might reflect variations of individual's energy intake levels (Willett et al., 1997). Total energy intake was calculated in our reference data-set from food frequency questionnaire using data from the British food composition tables (McCance and Widdowson, 1991).

ESFA was in the first instance unadjusted for effects of other food intakes, but in order to cope with confounding we also carried out an ESFA adjusting for effects of foods other than the index food.

So, let's assume that for the $i^{\text {th }}$ individual $V_{i j}$ is a vector of covariates which represent our confounders and $\zeta_{j}$ is a vector of regression coefficients for these covariates, our logistic model in equation 4.3.1.2 could take the general form,
$\log \left[\frac{E\left(Y_{i} / X_{i j}\right)}{1-E\left(Y_{i} / X_{i j}\right)}\right]=\beta_{o j}+\beta_{1 j} X_{i j}+\beta_{2 j}(\text { totalenergy })_{i}+\sum_{n=1}^{m} \beta_{m} V_{i m} \quad i=1 \ldots n$

The confounders $V_{i j}$ are actually considering here other foods. Because in practice we don't want to have to include all the foods in the regression equation (4.3.1.2), we need practical strategies to adjust for all the foods we're not interested in. Hence three different strategies of adjustment were employed in our logistic model in (4.3.1.3) as presented below.

### 4.3.2 Description of ESFA procedure adjusted for the first five principal components of diet

Firstly, the simulated dietary data were subjected to a PCA using a varimax rotation. Five principal components were identified and their principal components scores were calculated for each individual (see paragraph 2.4). An ESFA procedure adjusted for the resulting five rotated principal component scores (covariates) was carried out with the use of the logistic model in equation 4.3.1.3. Because net confounding by correlated foods and patterns could result to biased associations between diet and disease, adjustment for dietary patterns has been suggested by the literature (Imamura et al., 2009). We choose five components of diet because that was the number of dietary patterns of diet that were identified in the original population (see paragraphs 4.9.2 and 4.9.3).

### 4.3.3 Description of ESFA procedure adjusted for all the food intakes which were significant in the unadjusted ESFA) controlling the false discovery rate at $\mathbf{2 0 \%}$ level

In this case, as a first step we run an ESFA procedure adjusting for total energy intake as described in paragraph 4.3.1, keeping the food variables (covariates) with those regression coefficient estimates that had a P-value lower than a specific threshold. The threshold was determined by the Benjamini and Hochberg procedure controlling the rate of our false discoveries at $20 \%$ (see paragraph 4.4). Then we re-run (second step) an ESFA procedure adjusting for energy intake and for all these foods (covariates) that were significant in the first round of analysis (unadjusted ESFA) with the use of the logistic model in equation 4.3.1.3. This method is conceptually similar with the Iterative sure independence screening method (ISIS) proposed by Fan and Lv (2008). However in our case we choose our covariates based on a multiple test procedure (Benjamini and Hochberg) and not on a penalized likelihood method. Furthermore we aim to control the false discovery rate, whereas sure screening method focuses on missed discoveries (Fan \& Lv, 2008).

### 4.3.4 Description of ESFA procedure adjusted for a propensity score predicting the index food intake from the other food intakes

Propensity scores where first established by Rosenbaum and Rubin (1983) (Rosenbaum \& Rubin D. P., 1983) and can be defined as the probability of exposure to a specific treatment for randomised controlled trials or to a potential risk factor for observational studies given
observed covariates. The development of the score was first developed for binary exposures (Rubin, 1997) but it was extended and generalized for continuous and ordinal exposures (Kosuke Imai \& David A van Dyk., 2004). A generalized linear model could be employed for calculating the propensity score given an observed set of covariates. Once the propensity score is estimated it could be used as a confounder in a conventional multivariate outcome model, (Sturmer et al., 2006).

Specifically, in our simulation study we defined the propensity score $p s_{i j}$ as the predicted value of an individual's intake of a specific food $i$ given all the other intakes of foods that an individual consumed in our study with the use of a linear regression model. So let's assume that $X_{i j}$ is the $j^{\text {th }}$ food item for the $i^{\text {th }}$ individual of our sample of $n$ individuals, and $\beta_{j}$ are the regression coefficient for the $j^{\text {th }}$ food item then

$$
\begin{equation*}
p s_{i j}=\beta_{o}+\sum_{\substack{i=1 \\ i \neq j}}^{k} \beta_{j} X_{i j} \tag{4.3.4.1}
\end{equation*}
$$

The resulting regression model produces for each food item $j$ a corresponding propensity score according to 4.3.4.1. These propensity scores were considered as a confounder (covariate) in our ESFA procedure for each food item $j$ with the use of equation 4.3.1.3.

Before discussing about the "environment" that we want to evaluate our procedures (PCR, ESFA and adjusted ESFA) we will give a description of the multiple test procedures which are a vital part of our simulation study.

### 4.4 Multiple Test Procedures

### 4.4.1 Overview

In hypothesis testing, when we want to test the association between two variables in our sample, we are testing the null hypothesis (that is the two variables are unrelated and that any apparent difference in our sample is due to chance) over the alternative. A p-value is the probability of obtaining a test statistic for exploring this association at least as large as the one observed in our sample, assuming that the null hypothesis in our population is true. For example, when we want to test if the effect of a food intake variable on disease in our sample is zero, we observe a value of a z -test. P -value gives the probability that this z -value value to be as large as the one we observe, assuming that the population of z -values is following a standard
normal distribution. If this $p$-value is lower than an arbitrary cut-off value of significance $\alpha$, which is our type I error, then our effect estimate is not zero.

However, as Rothman (1990) points out "if $n$ independent associations are examined for statistical significance, the probability that at least one of them will be found statistically significant is $1-(1-\alpha)^{\mathrm{n}}$, if all $n$ of the individual hypotheses are true". So for example, in our case, when an ESFA is applied in our data-set 1 (see paragraph 4.8.1) 217 separate logistic regression models are employed to explore the association between each separated food item and disease, and consequently 217 hypotheses are tested. So if we assume that our type I error is 0.05 for the 217 independent hypothesis, then the probability that at least one of them will be found statistically significant is $1-(1-\alpha)^{217}=0.999$, assuming that all of 217 of the null hypotheses are true. So we have a probability of almost 1 to report a false positive result, if all of the null hypotheses are true.

Multiple test procedures are used for tackling the issue of multiple comparisons. Two wellestablished methods are presented to this thesis as a response to this problem; Bonferroni inequality and the Benjamin and Hochberg method of the rate of false discoveries which we will describe below.

### 4.4.2 Bonferroni inequality and family-wise error rate (FWER)

Let's assume that $\left\{P_{1}, P_{2} \ldots P_{m}\right\}$ is a set of our observed $p$-values for our $m$ null hypotheses $\left\{\mathrm{H}_{1}, \mathrm{H}_{2} \ldots \mathrm{H}_{\mathrm{m}}\right\}$ and $\alpha$ is our acceptable type I error. Bonferroni inequality is defined by the formula
$\operatorname{Pr}\left[\min \left(\mathrm{P}_{\mathrm{j}}: 1 \leq \mathrm{j} \leq \mathrm{m}\right) \leq \alpha / \mathrm{m}=\mathrm{q}\right] \leq \alpha \quad(4.4 .2 .1)$
or if P-values are ordered in ascending order $\mathrm{P}_{(1)} \leq \mathrm{P}_{(2)} \leq \ldots \leq \mathrm{P}_{(\mathrm{m})}$, then
let $k$ to be the largest $(i)$ for which $P_{(i)}<\frac{a}{m}$
then reject all the corresponding $\mathrm{H}_{(\mathrm{i})}, \mathrm{i}=1,2, \ldots, \mathrm{k}$.

In the Bonferroni procedure this upper bound $\mathrm{q}=\alpha / \mathrm{m}$ is denoted as the family-wise error rate (FWER) and defined as the probability that at least one of the null hypotheses in our set will be rejected. More formally, If the upper bound $\mathrm{q}(0<\mathrm{q}<1)$ is a FWER then we can be $100(1-\alpha)$ $\%$ confident that all null hypothesis in the subset of $p$-values that are below $q$ are false
(Newson, 2003). However, controlling the FWER with the Bonferroni inequality is a conservative method, which could omit important associations between diet and disease, and could lead to false negative findings because of loss of power in our study. Furthermore, in some situations this could be unnecessary. For example, FWER is important when a conclusion from the various individual inferences is likely to be erroneous when at least one of them is.

In our case, if we want to test the hypothesis if each one of individual food item on the FFQ is separately associated with disease, we don't want to falsely accept a specific null hypothesis (that there is no association between a specific food item and disease) because some of the null hypothesis (for the other food items) are falsely rejected (Benjamini and Hochberg 1995). In order to tackle with these problems Benjamini and Hochberg suggested a less conservative procedure the maximum permissible false discovery rate (FDR).

### 4.4.3 Benjamini and Hochberg procedure and false discovery rate (FDR)

As defined by Benjamini and Hochberg (1995), false discovery rate (FDR) is the expected proportion of true null hypothesis that declared significant (false positives) among all the hypotheses that declared significant and controls the proportion of the rejected null hypotheses which are erroneously rejected.

Let's assume that $\left\{\mathrm{P}_{1}, \mathrm{P}_{2} \ldots \mathrm{P}_{\mathrm{m}}\right\}$ is a set of our observed p -values for the corresponding m null hypotheses $\left\{\mathrm{H}_{1}, \mathrm{H}_{2} \ldots \mathrm{H}_{\mathrm{m}}\right\}$ and $\mathrm{q}^{*}$ is the maximum permissible FDR. P-values are ordered in ascending order $\mathrm{P}_{(1)} \leq \mathrm{P}_{(2)} \leq \ldots \leq \mathrm{P}_{(\mathrm{m})}$.
let $k$ to be the largest $i$ for which $\mathrm{P}_{(\mathrm{i})}$ for which $P_{(i)} \leq \frac{i}{m} q^{*}$
then reject all the corresponding $\mathrm{H}_{(\mathrm{i})}, \mathrm{i}=1,2, \ldots, \mathrm{k}$.

Benjamini \& Hochberg (1995) proved under general conditions that the above procedure for rejecting hypotheses leads to a false discovery rate of no more than $q^{*}$. If all of null hypothesis are true then the FDR is FWER.

### 4.5 Evaluation of performance of ESFA and PCR procedures in each simulation experiment

First, we investigated the statistical power with which ESFA and PCA could detect whether there was any association between diet and disease. For ESFA, using the methodology described in Chapter 4.3.1 we considered that an association had been found if any of the food intakes were statistically significantly associated with disease after applying a Bonferroni correction (see chapter 4.4.2) for the number of foods (family-wise $\mathrm{P}<0.05$ ) (Miller, 1981). For PCR, using the methodology described in paragraph 4.3.1, we considered an association had been found if any of the dietary patterns were significantly associated with disease after applying a Bonferroni correction (see Chapter 4.4.2) for the number of patterns identified. In both situations, we control our results for the FWER because we wanted to be $95 \%$ confident that all of statistical significant associations are real.

We also wanted to see how well the two procedures identified the specific combinations of foods that were causally linked with disease. We compared the power and the false discovery rate (FDR) of ESFA and PCA for detecting these associations. In this context we extend the concept of "power" to mean the proportion of foods included in the model which were identified as significant. The FDR is the proportion of discoveries, or significant findings, which are false (Figure 4.5.2.1). More formally, power is defined as the number of true significant results identified by the method (True Positives) divided by the number of foods that are causally linked with disease (TP/ (FN+TP)), and false discovery rate (FDR) is defined as the number of false significant results (False Positives) identified by the method divided by the total number of significant results identified by the method (FP / (FP +TP )), or 0 if FP + $\mathrm{TP}=0$ ) .

Figure 4.5.2.1. How the results of dietary analyses can be broken down

|  | Foods not causally <br> linked with disease | Foods causally <br> linked with disease |
| :--- | :--- | :--- |
| Foods declared <br> non-significant | True Negatives <br> (TN) | False Negatives <br> (FN) |
| Foods declared <br> significant | False Positives <br> (FP) | True Positives (TP) |

Power $=T P /(F N+T P)$

False Discovery Rate $\quad=\mathrm{FP} /(\mathrm{FP}+\mathrm{TP})$ if FP + TP $>0$
$=0 \quad i f \mathrm{FP}+\mathrm{TP}=0$

For ESFA, we considered there was a "significant" effect of a food if it was identified as such using the multiple testing procedure of Benjamini and Hochberg, with a nominal false discovery rate set to $20 \%$ (Benjamini \& Hochberg, 1995) (see chapter 4.4.3). For PCA, we considered there was a "significant" effect of a food if it had correlation $>0.3$ or $<-0.3$ with a dietary pattern that was significantly associated with disease ( $\mathrm{P}<.0 .05$ ) - this being the way in which individual foods tend to be highlighted in a PCA (median value of all the studies in our systematic review was 0.3 - see paragraph 3.2.5). Furthermore, we control the rate of our false discoveries at $20 \%$ since we want to be $80 \%$ confident that some of the statistical significant associations that we observe are real, or that $80 \%$ of these associations are expected to be real. Note that the Benjamini-Hochberg procedure is designed to control the FDR at no more than the nominal level, but here false discoveries (of foods) occur not just as random errors, but also because of confounding with other foods, so the nominal rate may be exceeded.

### 4.6 Standards error of our Monte Carlo simulations

In each simulation, we are sampling with replacement N individuals from our real-reference data-set (Data-set 1-see paragraph 4.8.1, Data-set 2 -see paragraph 4.8.2) and create a sample. In each sample we construct our diet ( $X_{b}^{*}$ ) and disease $\left(Y_{b}^{*}\right)$ variables according to paragraph 4.2, and we estimate our three parameters of interest according to paragraphs 4.3, 4.4 and 4.5; (i) power ( $\hat{\theta}_{1}^{*}$ ) with which ESFA and PCA could detect whether there was any association between diet and disease; (ii) power ( $\hat{\theta}_{2}^{*}$ ) with which ESFA and PCA could detect specific combination of foods that are causally linked to disease and (iii) and false discovery rate ( $\hat{\theta}_{3}^{*}$ )
of ESFA and PCA for detecting these combinations of foods. We want to calculate the corresponding standard errors for these estimation results.

Thus, in case (i) our estimate of power ( $\hat{\theta}_{1}^{*}$ ) is the proportion of statistical significant associations that were declared significant out of $B$ replications (in our case $B=10000$ ). Standard error of this proportion is calculated by the formula
$s \hat{e}\left(\hat{\theta}_{1}\right)=\sqrt{\frac{1}{B} p(1-p)}$

In addition, we construct our standard errors for our estimates (ii) power ( $\hat{\theta}_{2}^{*}$ ) and (iii) false discovery rate $\left(\hat{\theta}_{3}^{*}\right)$ according to the following algorithm (Efron \& Tibshirani, 1993).

1. Select B independent samples $\left(X_{b}^{*}, Y_{b}^{*}\right)$ of size n from our reference data-set. In our case $\mathrm{B}=10000$ replications
2. We estimate our two parameters of interest for each sample

$$
\hat{\theta}_{c b}^{*}, \text { for } b=1, \ldots, B \text { and } \mathrm{c}=2,3
$$

3. we estimate the standard error $\operatorname{se}\left(\hat{\theta}_{c}\right)$ by the sample standard deviation and according to the formula
$s \hat{e}\left(\hat{\theta}_{c}\right)=\sqrt{\frac{1}{B-1} \sum_{b=1}^{B}\left(\hat{\theta}_{c b}^{*}-\hat{\theta}_{c}^{*}\right)^{2}}, \mathrm{c}=2,3$
where $\hat{\theta}_{c}^{*}=\frac{1}{B} \sum_{b=1}^{B}\left(\hat{\theta}_{c b}^{*}\right)$

### 4.7 Sample size calculation of our simulation experiments

Using the "powerlog" sample size calculation routine in Stata (Ender, 2002), we determined that a sample size of 330 would achieve $80 \%$ power at the $5 \%$ significance level to detect an odds ratio of 1.5 per standard deviation, using an unadjusted logistic regression with no allowance for multiple testing(Table 4.7.1).

Table 4.7.1.Power Analysis

| Power | $\mathbf{N}$ |
| :--- | :--- |
| 0.60 | 194 |
| 0.65 | 221 |
| 0.70 | 252 |
| 0.75 | 287 |
| 0.80 | 328 |
| 0.85 | 380 |
| 0.90 | 451 |

### 4.8 Reference data sets

Our two real dietary data-sets and source of the food correlation matrices for our simulations were comprised of adults living in Greenwich as part of the F.L.A.G survey and Ipswich and Norwich as part of the UK ECRHS II diet survey. We used only controls from both data-sets in order to have a more representative reference correlation matrix for our simulation process. Food frequencies were converted to intakes in g/d by multiplying frequency of consumption by the weight of standard portion sizes using British food composition tables (Paul, Southgate \& Buss, 1986) and standardised to have a mean 0 and a standard deviation of 1.

### 4.8.1 Food, Lifestyle \& Asthma in Greenwich Survey (F.L.A.G)

The original dataset being used in our study was based on 856 adults aged 16-50 years without asthma who responded to an asthma survey in a random sample of adults 16-50, registered with 40 general practices in Greenwich, South London, UK, in autumn 1996 (Marks et al., 1997, Premaratne et al., 1999). Individuals were mailed a dietary questionnaire in September 1997. Usual diet was assessed (previous 12 months) using a food frequency questionnaire (FFQ) based on one used previously (Calvert et al., 1997). Food frequency questionnaires (FFQ) recorded a consumption as frequencies of 217 different foods (from never to 6d a week) and drinks.

### 4.8.2 European Community Respiratory Health Survey II (ECRHS II UK)

ECRHS-I ran from 1990 to 1995 . At each centre, a random sample of at least 3000 adults aged 20-44 years was selected using a local sampling frame. From those who responded, a random sample of at least 600 adults was selected to undergo a detailed clinical examination. Eight to ten years later, these subjects were contacted to take part in a follow-up study
(ECRHS-II) and invited to a local clinic for further assessments, including an intervieweradministered questionnaire (European Community Respiratory Health Survey II Steering Committee, 2002).

Dietary assessments were included in ECRHS-II at some centres, though the method and protocol differed between countries. In the present study, we report results from 2 centres in UK, where FFQ were administered: Ipswich and Norwich. Three thousand three hundred and eighty-seven adults at these centres were contacted to take part in ECRHS-II. The UK FFQ was adapted from one developed for EPIC-UK (Bohlscheid-Thomas et al., 1997). It recorded a consumption of 198 different foods over the last 12 months as frequencies (from never to 7 d a week) and number of portions consumed on each of these days (portions being defined on the questionnaire). Some aggregation of food items into food groups was performed and this process led to a list of 74 food groups whose intake in $\mathrm{g} / \mathrm{d}$. In our study, we included 201 adults aged 29-54 years living in Ipswich and Norwich.

### 4.9 Construction of a simplified "Western" dietary pattern

### 4.9.1 Overview

Schulze et al 2003 (Schulze et al., 2003) proposed a method to construct a simpler form of a dietary pattern variable applied previously in the field of psychology (Comrey, 1988) and described extensively by Jolliffe (Jolliffe, 2010). The main idea of this method is to associate food items with each of the first few dietary patterns and then retaining those food items which are more strongly associated with the first dietary patterns. The choice of which food items to retain for each pattern should be determined by looking at the strength of the relationship between the food item and the dietary pattern So, in our case, a simplified "Western" pattern was constructed by selecting 30 items from data-set 1 (see paragraph 4.9.2), and the 10 items from data-set 2 (see paragraph 4.9.3), that were most strongly positively correlated with the Western principal component of the population.

### 4.9.2 Simplified "Western" Pattern derived from the F.L.A.G data-set

Dietary patterns were identified with the use of PCA from the F.L.A.G data-set. The principal components were rotated (varimax rotation) and the number of patterns were determined by examination of the scree plot of the eigenvalues. We extracted five components (dietary patterns), which explained $17 \%$ of the variance in the original 216 items. Individual foods that correlated $>0.3$ or $<-0.3$ with the varimax rotated principal components labelled the
dietary pattern. More detailed information of the analysis is presented elsewhere (Bakolis et al., 2010). From the five patterns that we identified we randomly choose the one that was labelled as "Western". The following 30 food items or groups where highly positively correlated with the "Western" pattern and consisted our newly constructed simplified "Western" pattern are presented in table 4.9.2.1. This newly constructed simplified "Western" dietary pattern was strongly correlated with the original western pattern of the F.L.A.G dataset ( $r=0.84$ ).

Table 4.9.2.1. List of foods comprising a "Western" pattern for each data-set ${ }^{\text {a }}$

| Data-set 1 | Data-set 2 |
| :--- | :--- |
| roast potatoes | sausages |
| ham | donuts, pastries and tarts |
| ice cream | beer |
| pork - roast, chops | corned beef and luncheon meat |
| pork stew, casserole | hard cheeses |
| omelette/scrambled egg | tomato ketchup |
| fruit pies, tarts, crumbles | pizza |
| beef stew,casserole,mince,curry | beef burger |
| sponge cakes | fried egg, scrambled egg, omelette |
| fried fish in batter/breadcrumb | chips |
| baked beans |  |
| chocolate biscuits |  |
| sandwich/cream biscuits |  |
| corned beef, spam, luncheon meat |  |
| white bread and rolls |  |
| fizzy soft drinks e.g. coke |  |
| bacon |  |
| fried egg |  |
| milk chocolate |  |
| bread crumbed chicken e.g. chicken nuggets |  |
| crisps |  |
| sponge puddings |  |
| tomato ketchup |  |
| chocolate snack bars |  |
| meat pizza |  |
| other fried snacks |  |
| chips |  |
| sausages - beef, pork |  |
| beef burger, hamburger |  |
| pies/pasties/sausage rolls |  |
| a Data-set 1 is from the FLAG survey; Data-set 2 is from the UK ECRHS II survey |  |

### 4.9.3 Simplified "Western" Pattern derived from the UK ECRHS II data-set

Similarly, dietary patterns were identified with the use of PCA to the UK ECRHS II data-set. We extracted five principal components from the examination of the scree plot, which were rotated orthogonally (varimax rotation) for better interpretation. Components explained 24.7\% of the overall variation in the original 74 food items. Correlations between our 74 food items and our dietary patterns are shown in table 4.9.3.1. For consistency purposes, we constructed again a simplified "Western" dietary pattern from the following 10 food items that were positevely correlated highly with the "Western" pattern (Table 4.9.2.1). This newly constructed simplified "Western" dietary pattern was strongly correlated with the original "Western" pattern of the UK ECRHS II dataset $(r=0.85)$.

Table 4.9.3.1. Correlations between food intakes and each of the five orthogonal rotated dimensions of diet, only correlations $>\mathbf{0 . 3 0}$ and $<-0.30$ are included in the table.

| pattern |  |  |
| :---: | :---: | :---: |
| fruit and | meat and 2 |  |
| vegetarian vegetables | deserts and |  |
| vegetables | Western | cereals |

food items

| soy cheese, tofu, quern, grains | 0.66 | - | - | - | - |
| :--- | :---: | :---: | :---: | :---: | :---: |
| honey |  |  |  |  |  |
| lentils, dahl, mixed bean | 0.62 | - | - | - | - |
| casserole | 0.57 | - | - | - | - |
| vegetarian paste | 0.54 | - | - | - | - |
| rice and rice dishes | 0.53 | - | - | - | - |
| kiwi, mango and pineapple | 0.50 | - | - | -0.40 | - |
| garlic | 0.48 | - | - | - | -0.35 |
| peppers | 0.45 | - | - | - | - |
| tomato | 0.45 | - | - | - | - |
| bean sprouts | 0.38 | - | - | - | - |
| apple | - | 0.78 | - | - | - |
| pear | - | 0.75 | - | - | - |
| orange | - | 0.74 | - | - | - |
| banana | - | 0.64 | - | - | - |
| peach and nectarine | - | 0.62 | - | - | - |
| grapes | - | 0.57 | - | -0.39 | - |
| other fruit juice | 0.31 | - | - | - | - |
| sliced meat | - | - | 0.58 | - | - |
| beef steak |  |  |  |  |  |
| minced beef, meat stew and | - | - | 0.51 | - | - |
| casserole | - | - | 0.51 | - | - |
| sausages | - | - | 0.50 | 0.31 | - |
| liver | - | - | 0.49 | - | - |


| pork chops | - | - | 0.43 | - | - |
| :---: | :---: | :---: | :---: | :---: | :---: |
| broccoli, cabbage and |  |  |  |  |  |
| cauliflower | 0.37 | 0.33 | 0.41 | - | - |
| carrots | 0.37 | - | 0.41 | - | - |
| bacon | - | - | 0.37 | - | - |
| potato - boiled / mashed / baked | - | - | 0.37 | - | - |
| boiled egg | - | - | 0.35 | - | - |
| peas | - | - | 0.35 | - | - |
| green beans | 0.36 | 0.34 | 0.32 | - | - |
| poultry | - | - | 0.32 | - | - |
| pate | - | - | 0.32 | - | - |
| chips | - | - | - | 0.55 | - |
| fried egg, scrambled egg, omelette | - | - | - | 0.50 | - |
| beef burger | - | - | - | 0.48 | - |
| pizza | - | - | - | 0.40 | - |
| hard cheeses | - | - | - | 0.37 | 0.34 |
| tomato ketchup | - | - | - | 0.37 | - |
| corned beef and luncheon meat | - | - | - | 0.36 | - |
| beer | - | - | - | 0.34 | - |
| donuts, pastries and tarts | - | - | - | 0.31 | 0.31 |
| herbal tea | - | - | - | -0.32 | - |
| yoghurt | - | - | - | -0.38 | - |
| raspberries, red currants, blackcurrants | - | - | - | -0.46 | - |
| cakes, puddings and desserts | - | - | - | - | 0.55 |
| breakfast cereals | - | - | - | - | 0.44 |
| chocolate | - | - | - | - | 0.44 |
| milk and milky drinks | - | - | - | - | 0.44 |
| choc bars and cereal bars | - | - | - | - | 0.39 |
| bread and rolls | - | - | - | - | 0.37 |
| ice cream | - | - | - | - | 0.36 |
| butter | - | - | - | - | 0.35 |
| wine | - | - | - | - | -0.35 |
| jam and marmalade | - | - | - | - | - |
| peanut butter and choc spreads | - | - | - | - | - |
| biscuits | - | - | - | - | 0.65 |
| cream cheese | - | - | - | - | - |
| cottage cheese | - | - | - | - | - |
| soft cheeses | - | - | - | - | - |
| quiche | - | - | - | - | - |
| fish fillets / cakes / fingers | - | - | - | - | - |
| tinned fish | - | - | - | - | - |
| soup | - | - | - | - | - |
| strawberries | - | - | - | - | - |
| tinned or stewed fruit | - | - | - | - | - |


| nuts | - | - | - | - | - |
| :--- | :--- | :--- | :--- | :--- | :--- |
| orange juice | - | - | - | - | - |
| fizzy drinks | - | - | - | - | - |
| tea - black and green | - | - | - | - | - |
| coffee (not decaffeinated) | - | - | - | - | - |
| decaffeinated coffee | - | - | - | - | - |
| cider | - | - | - | - | - |
| fortified wine | - | - | - | - | - |
| liqueurs and spirits | - | - | - | - | - |

### 4.10 Specification of simulation parameter values and analysis of simulation experiments

For each simulated experiment a number of 4, 10 (UK ECHRS II data-set) and 10, 30 (F.L.A.G data-set) food items with the same effect are chosen randomly and are causally associated with disease in 3 different ways. Additionally a simplified "Western" dietary pattern from our two datasets (F.L.A.G and E.C.H.R.S II) is causally associated with disease in 2 different ways. The number of individuals was set to be 100 (in only two case), 300, 600, 1200, 2400 and 4800 and the number of varimax rotated principal components was chosen to be 2, 5 and 10. Comparisons were made between our ESFA and PCA procedure for each different combination of our parameter values. ESFA procedure was further adjusted for effect of other foods in 3 different ways. Average percentages of power and false discovery rate were calculated, and after looking on the results of 10000 simulation trials, we observed that the standard errors for these percentages were $<0.5 \%$.

### 4.11 Null simulations

Null simulations were conducted for testing the validity of our programming work. In this case no randomly selected number of foods or a simplified "Western" pattern was causally linked with disease (Model 0). So in this case we assumed a logistic model for the disease risk, $p$ that is

$$
p=\frac{1}{1+e^{-0.15}},(4.11 .1)
$$

Power for detecting any association between diet and disease was controlled at the desired level of $5 \%$ (Table 4.11.1).

Table 4.11.1. Power (\%) of exhaustive single food analysis (ESFA) and principal components analysis (PCA) to detect any association between diet and disease (all estimates of power have standard error $<\mathbf{0 . 5 \%}$ ).

| Data-set and <br> model $^{\text {a }}$ | Sample Size | ESFA |  | PCA |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
|  |  |  | Number of Components |  |  |
|  |  | Power | Power | Power | Power |
|  |  |  |  |  |  |
| Data-set 1 |  | 1.0 | 4.3 | 3.0 | 2.7 |
| Model 0 | 300 | 2.0 | 4.9 | 3.7 | 2.4 |
|  | 600 | 3.3 | 4.1 | 4.0 | 3.7 |
|  | 1200 | 3.7 | 4.8 | 4.6 | 4.2 |
|  | 2400 | 4.2 | 4.6 | 4.9 | 4.8 |
| Data-set 2 |  |  |  |  |  |
| Model 0 | 3800 | 2.3 | 4.5 | 3.3 | 3.6 |
|  | 300 | 3.0 | 4.7 | 4.6 | 4.1 |
|  |  | 3.8 | 4.2 | 4.1 | 4.6 |
|  | 1200 | 4.9 | 4.6 | 4.5 | 4.6 |
|  | 2400 | 4800 | 4.5 | 4.6 | 4.9 |

${ }^{\text {a }}$ Data-set 1 is from the FLAG survey (Shaheen et al., 2001); Data-set 2 is from the UK ECRHS II survey (Hooper et al., 2010).In Model 0 no foods or simplified "Western" pattern are selected at random in each replication from the foods on the FFQ.

The majority of null simulations in Table 4.11.1 don't reach their nominal level. Power of PCA is around $5 \%$ only for 2 and 5 principal components. Power of ESFA and PCA for 10 components reaches the nominal level only for large sample sizes (2400,4800). In addition, as the number of components increases power of PCA descreases and reaches a nominal level much lower than $5 \%$ for small sample sizes $(300,600,1200)$.

One potential explanation is that the Bonferroni correction is a conservative method, meaning that the actual family-wise error rate will be smaller than the value it is fixed at when results are correlated. Suppose, for example, we are applying the Bonferroni correction to the results of 100 tests. For any result to be significant it must have a $\mathrm{P}<0.05 / 100$. If all the tests are independent, then the probability under the global null hypothesis that none of them are significant is $(1-0.05 / 100)^{100} \cong 0.95$. So, the probability that at least one result is significant under the null hypothesis is 0.05 . However, in the extreme example of dependence, if we repeat the same test on the same data 100 times, the probability would have to be less than $0.05 / 100=0.0005$ which is the probability that at least one result is significant under the null hypothesis. So, depending on how correlated the P -values are, the family-wise error rate might
be lower. So, due to highly correlated dietary intake variables in our simulation datasets, we don't expect the values of power under the null model in our simulations to be $5 \%$, but lower than this value. This could be observed from Table 4.11.2, where moderate correlation coefficient values are observed between the components for five and ten principal components but not for two.

Furthermore, we observed for ESFA that power increased with sample size. However this pattern was not observed so clearly for PCA. One potential explanation that power increases with sample size is due to the effect of skewness of our dietary intake data. As we can see from Figure 4.11 .1 , dietary pattern intake variables derived from the F.L.A.G study are normally distributed but food intake variables are highly skewed to the right.

In the context of hypothesis testing this translates into true tail probabilities that are higher than nominal in the upper tail, resulting in fewer rejections, than there would be under a normal parent for one-tailed tests (Boos et al.,1998, Rieneke et al., 2003). In a simulation study (Rieneke et al., 2003) where the type I error rates were calculated, a strong dependence was shown between the error rates, skewness and sample size of the data according to the formula $\sqrt{\beta}(X) / \sqrt{n}$, where $n$ is the size of the sample and $\sqrt{\beta}(X)$ is the measure of skewness. Particularly, when skewness was close to zero error rates were close to what we would have expected even for small sample sizes. However if the skewness is far from zero sample size must be large in order for error rates to be close to what we should have expected (Reineke et al., 2003). Thus, because of the skewness of our food intake data we have fewer rejections of the null hypothesis for small sample sizes, and as our sample size increases our power estimations are closer to the desirable error rate of $5 \%$. The implications of lower significance level for ESFA in our null simulations could potentially lead to an underestimation of the average percentages of power of ESFA for small sample sizes in our monte carlo simulations reported in Chapter 5.

Table 4.11.2. Correlation coefficient values between randomly identified dietary patterns for sample size of $\mathbf{6 0 0}$ from the simulated F.L.A.G survey data-set. Different numbers of dietary patterns were derived with the use of PCA ( 2,5 and 10 ).

| 2 Randomly simulated Principal Components |  |  |  |  |  |  |  |  |  |  |
| :--- | ---: | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Dietary pattern | I | II |  |  |  |  |  |  |  |  |
| I | 1.00 |  |  |  |  |  |  |  |  |  |
| II | -0.03 | 1.00 |  |  |  |  |  |  |  |  |
| 5 Randomly simulated Principal Components |  |  |  |  |  |  |  |  |  |  |
| Dietary pattern | 1 | 2 | 3 | 4 | 5 |  |  |  |  |  |
| 1 | 1.00 |  |  |  |  |  |  |  |  |  |
| 2 | -0.16 | 1.00 |  |  |  |  |  |  |  |  |
| 3 | 0.21 | 0.12 | 1.00 |  |  |  |  |  |  |  |
| 4 | 0.11 | -0.11 | 0.07 | 1.00 |  |  |  |  |  |  |
| 5 | 0.15 | -0.03 | 0.10 | 0.03 | 1.00 |  |  |  |  |  |
| $\mathbf{1 0}$ randomly simulated principal components |  |  |  |  |  |  |  |  |  |  |
| Dietary pattern | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 1 | 1.00 |  |  |  |  |  |  |  |  |  |
| 2 | 0.35 | 1.00 |  |  |  |  |  |  |  |  |
| 3 | -0.21 | -0.20 | 1.00 |  |  |  |  |  |  |  |
| 4 | 0.16 | 0.04 | 0.04 | 1.00 |  |  |  |  |  |  |
| 5 | -0.06 | -0.03 | 0.19 | -0.05 | 1.00 |  |  |  |  |  |
| 6 | 0.14 | 0.06 | 0.14 | 0.17 | 0.12 | 1.00 |  |  |  |  |
| 7 | 0.19 | 0.09 | -0.03 | 0.14 | 0.01 | 0.12 | 1.00 |  |  |  |
| 8 | 0.14 | 0.16 | -0.14 | -0.03 | -0.06 | -0.04 | 0.03 | 1.00 |  |  |
| 9 | 0.17 | 0.19 | -0.04 | 0.04 | 0.05 | 0.08 | 0.00 | 0.03 | 1.00 |  |
| 10 | -0.19 | -0.12 | 0.16 | -0.07 | 0.12 | 0.03 | -0.05 | -0.10 | -0.03 | 1.00 |

Figure 4.11.1. Histograms of random selected food intake and 4 randomly selected principal component analysis.


### 4.12 Programming

In this project, Monte Carlo simulations were developed and programmed using Stata 10 (Stata Corporation, College Station, Texas USA). Simulation programs were created from scratch; one for selecting randomly combinations of foods; two for generating a simplified "Western" dietary pattern; one for examining the power with which ESFA and PCR could detect whether there was any association between diet and disease; one for examining the power and FDR with which ESFA and PCR could detect a specific combination of foods; and three for examining for examining the power and FDR with which ESFA could detect a specific combination of foods under different ways of adjustment. Programs were designed to allow for a range of different number of

- bootstrap replications
- size of bootstrap samples
- size of baseline risks
- Effect sizes of the randomly selected number of food items and of the simplified "Western" patterns.
- size of the number of foods that are causally associated with disease
- ways that number of foods are causally associated with disease in the simulation model
- ways that a simplified "Western" pattern was causally associated with disease in the simulation model
- cut-off points of correlation coefficients of food items with the rotated dietary pattern
- cut-off points of family wise error rate (FWER) and false discovery rate (FDR)
- Number of rotated principal components identified.

Computer algorithms and commands being used are presented at the Appendix II

## 5 Results

### 5.1 Introduction

5.2 Power with which ESFA and PCA could detect whether there was any association between diet and disease
5.3 Power and FDR with which ESFA and PCA could identify specific combinations of foods between diet and disease
5.4 Power and FDR with which ESFA and PCA could identify specific combinations of foods between diet and disease adjusted in three different ways

### 5.1 Introduction

In this chapter results from our simulations are presented. Paragraph 5.2 present average estimates of percentages of power and paragraph 5.3 presents average estimates of percentages of power and FDR for different sample sizes and different number of principal component scenarios. Paragraph 5.4 shows average percentages of power and false discovery rates for different ways of adjustment of ESFA method for different sample sizes.

### 5.2 Power with which ESFA and PCA could detect whether there was any association between diet and disease

Tables 5.2.1, 5.2.2, 5.2.3 display average estimates of power of exhaustive single food analysis (ESFA) and principal component analysis (PCA), when 1 in 7 or 1 in 20 foods are randomly selected and causally linked with disease. Additionally, PCA and ESFA are evaluated when a simplified "Western" pattern is causally associated with disease.

Both methods had considerable power to detect any statistical effect between diet and disease. In the majority of the scenarios we investigated, ESFA had greater power than PCA to detect an association between diet and disease, when randomly selected foods where causally associated with disease Although PCA performed slightly better in some (21/240) simulations this was only when small sample sizes of 300 and 600 were used. The power of ESFA was increased with sample size. Average percentages of power estimates of ESFA didn't present a symmetrical pattern between model 1 (all selected food intakes are negatively associated with disease) and model 3 (all selected food intakes are positively associated with disease), and we get intermediate power estimates for Model 2 (half the selected food intakes are negatively associated and half are positively associated with disease) (Table 5.2.1 and 5.2.2). Similar qualitatively simulation results are observed in Table 5.2.3 for model 4 and 5 when a "Western" dietary pattern is causally associated with disease.

PCA outperforms ESFA in all scenarios when a "Western" pattern is causally associated with disease for sample sizes lower than 600 for all different odds ratio values (1.1, 1.3, and 1.5). The power of principal component analysis increases with the number of principal components and sometimes decreases. Furthermore, for PCA, averages percentages of power to detect any effect of diet on disease is roughly the same when all effects of foods are positive as when all effects are negative, and the power is less when there is a mixture of positive and negative effects (only for ECHRS II survey dataset at Table 5.2.2 we observe
intermediate power when there is a mixture of positive and negative effects of foods on disease).

Table 5.2.1. Power (\%) of exhaustive single food analysis (ESFA) and principal components analysis (PCA) to detect any association between diet and disease (all estimates of power have standard error $\mathbf{< 0 . 5 \%}$ ). In models $\mathbf{1 - 3}$, one in seven foods ( $\mathbf{3 0}$ in data-set 1 and 10 in data-set 2) are selected at random in each replication from the foods on the FFQ. Family-wise error of Bonferroni correction at 5\%.

| Data-set and model ${ }^{\text {a }}$ | Average Number of Cases | Sample Size | ESFA | PCA |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Number of components |  |  |
|  |  |  |  | 2 | 5 | 10 |
| Data-set 1 (FLAG) |  |  |  |  |  |  |
| Model 1 <br> (All foods negatively associated with disease) | 79.3 | 300 | 41.0 | $98.2{ }^{\text {b }}$ | $98.2{ }^{\text {b }}$ | $97.2{ }^{\text {b }}$ |
|  | 158.1 | 600 | 93.4 | $99.8{ }^{\text {b }}$ | $99.8{ }^{\text {b }}$ | $99.4{ }^{\text {b }}$ |
|  | 316.5 | 1200 | 100.0 | 100.0 | 100.0 | 100.0 |
|  | 634.0 | 2400 | 100.0 | 100.0 | 100.0 | 100.0 |
|  | 1266.8 | 4800 | 100.0 | 100.0 | 100.0 | 100.0 |
| Model 2 <br> (Half foods positively, half negatively associated with disease) | 81.9 | 300 | 77.2 | 38.4 | 44.8 | 48.4 |
|  | 163.8 | 600 | 99.2 | 56.0 | 69.7 | 80.0 |
|  | 327.8 | 1200 | 100.0 | 78.6 | 87.9 | 93.2 |
|  | 655.2 | 2400 | 100.0 | 86.8 | 95.2 | 100.0 |
|  | 1311.5 | 4800 | 100.0 | 93.7 | 100.0 | 100.0 |
| Model 3 <br> (All foods positively associated with disease) | 86.4 | 300 | 94.2 | $98.1{ }^{\text {b }}$ | $99.9{ }^{\text {b }}$ | $98.4{ }^{\text {b }}$ |
|  | 172.7 | 600 | 99.9 | 99.9 | 99.9 | 99.9 |
|  | 345.7 | 1200 | 100.0 | 100.0 | 100.0 | 100.0 |
|  | 690.6 | 2400 | 100.0 | 100.0 | 100.0 | 100.0 |
|  | 1381.0 | 4800 | 100.0 | 100.0 | 100.0 | 100.0 |

Data-set 2 (ECRHS
II)

| Model 1 | 59.8 | 300 | 30.2 | $44.5^{\mathrm{b}}$ | $49.5^{\mathrm{b}}$ | $49.4^{\mathrm{b}}$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| (All foods negatively | 119.4 | 600 | 82.1 | 66.2 | 77.8 | 82.0 |
| associated with disease) | 238.9 | 1200 | 99.7 | 82.8 | 92.5 | 96.6 |
|  | 478.0 | 2400 | 100.0 | 92.2 | 97.9 | 99.7 |
|  | 956.3 | 4800 | 100.0 | 96.1 | 99.5 | 99.9 |
| Model 2 | 62.3 | 300 | 65.1 | 42.3 | 53.6 | 58.9 |
| (Half foods positively, half | 124.7 | 600 | 97.7 | 58.3 | 77.9 | 87.3 |
| negatively associated with | 249.2 | 1200 | 99.9 | 74.3 | 91.3 | 97.9 |
| disease) | 498.4 | 2400 | 100 | 84.7 | 97.4 | 99.8 |
|  | 997.3 | 4800 | 100.0 | 92.3 | 99.4 | 99.9 |
| Model 3 | 65.1 | 300 | 81.2 | 61.6 | 71.4 | 76.3 |
| (All foods positively | 129.8 | 600 | 99.3 | 78.9 | 89.6 | 95.6 |
| associated with disease) | 260.1 | 1200 | 100.0 | 90.0 | 97.6 | 99.5 |
|  | 520.0 | 2400 | 100.0 | 95.2 | 99.3 | 99.9 |
|  | 1040.3 | 4800 | 100.0 | 97.8 | 99.8 | 99.9 |

[^0]Table 5.2.2. Power (\%) of exhaustive single food analysis (ESFA) and principal components analysis (PCA) to detect any association between diet and disease (all estimates of power have standard error $<\mathbf{0 . 5 \%}$ ). In models $\mathbf{1 - 3}$, one in twenty foods ( 10 in data-set 1 and 4 in data-set 2) are selected at random in each replication from the foods on the FFQ. Family-wise error of Bonferroni correction at 5\%.

| Data-set and model ${ }^{\text {a }}$ | Average Number of Cases | Sample <br> Size | ESFA | PCA |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Number of components |  |  |
|  |  |  |  | 2 | 5 | 10 |
| Data-set 1(FLAG) |  |  |  |  |  |  |
| Model 1 <br> (All foods negatively associated with disease) | 61.7 | 300 | 11.5 | $53.1{ }^{\text {b }}$ | $47.1{ }^{\text {b }}$ | $42.1{ }^{\text {b }}$ |
|  | 123.2 | 600 | 52.3 | $75.6{ }^{\text {b }}$ | $74.4{ }^{\text {b }}$ | $72.4{ }^{\text {b }}$ |
|  | 246.3 | 1200 | 95.5 | 89.3 | 90.7 | 91.7 |
|  | 492.4 | 2400 | 100.0 | 96.3 | 97.3 | 98.2 |
|  | 984.8 | 4800 | 100.0 | 99.7 | 99.3 | 99.9 |
| Model 2 <br> (Half foods positively, half negatively associated with disease) | 62.4 | 300 | 31.0 | 25.7 | 27.8 | 27.0 |
|  | 124.7 | 600 | 85.7 | 41.7 | 49.9 | 53.7 |
|  | 249.7 | 1200 | 99.8 | 60.2 | 70.6 | 80.3 |
|  | 499.5 | 2400 | 100.0 | 73.8 | 87.5 | 96.5 |
|  | 998.9 | 4800 | 100.0 | 85.5 | 96.5 | 99.6 |
| Model 3 <br> (All foods positively associated with disease) | 66.8 | 300 | 67.8 | 67.1 | 67.4 | 67.4 |
|  | 130.8 | 600 | 97.7 | 83.1 | 86.2 | 87.6 |
|  | 258.7 | 1200 | 99.9 | 92.8 | 95.9 | 97.4 |
|  | 510.5 | 2400 | 100.0 | 96.4 | 99.2 | 99.7 |
|  | 1010.7 | 4800 | 100.0 | 99.2 | 99.5 | 100.0 |

Data-set 2
(ECRHS II)

| Model 1 | 51.6 | 300 | 8.8 | $19.3^{b}$ | $16.5^{b}$ | $17.3^{b}$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| (All foods negatively | 103.3 | 600 | 38.8 | 32.4 | 38.6 | 36.7 |
| associated with disease) | 206.6 | 1200 | 89.6 | 50.9 | 64.0 | 68.4 |
|  | 413.2 | 2400 | 100.0 | 68.3 | 83.3 | 92.5 |
|  | 826.4 | 4800 | 100.0 | 82.3 | 94.5 | 98.9 |
| Model 2 | 52.3 | 300 | 43.1 | 23.9 | 30.1 | 32.4 |
| (Half foods positively, | 104.4 | 600 | 89.0 | 40.2 | 52.9 | 62.6 |
| half negatively | 209.0 | 1200 | 99.7 | 56.8 | 75.3 | 86.7 |
| associated with disease) | 418.6 | 2400 | 100.0 | 73.4 | 90.5 | 97.6 |
|  | 837.1 | 4800 | 100.0 | 84.5 | 97.6 | 99.9 |
| Model 3 | 54.3 | 300 | 60.4 | 35.9 | 43.9 | 47.8 |
| (All foods positively | 106.7 | 600 | 96.6 | 54.3 | 69.7 | 78.1 |
| associated with disease) | 219.3 | 1200 | 100 | 71.6 | 86.9 | 95.2 |
|  | 429.1 | 2400 | 100 | 84.3 | 96.1 | 99.4 |
|  | 847.5 | 4800 | 100 | 91.8 | 99.0 | 99.9 |

${ }^{\text {a }}$ Data-set 1 is from the FLAG survey (Shaheen et al., 2001); Data-set 2 is from the UK ECRHS II survey (Hooper et al., 2010). In Model 1 all selected food intakes are negatively associated with disease; in Model 2 half the selected food intakes are negatively associated and half are positively associated; and in Model 3 all selected food intakes are positively associated with disease. ${ }^{\mathrm{b}}$ Power of PCA exceeds that of ESFA.

Table 5.2.3. Power (\%) of exhaustive single food analysis (ESFA) and principal components analysis (PCA) to detect any association between diet and disease (all estimates of power have standard error $<\mathbf{0 . 5 \%}$ ). In model 4 and 5, foods being included in a simplified "Western" dietary pattern ( 30 in data-set 1 and 10 in data-set 2 - see paragraph 4.9.2 and 4.9.3) are used in each replication. Family-wise error of Bonferroni correction at 5\%.

| Data-set and model ${ }^{\text {a }}$ | Sample Size | ESFA | PCA |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Number of Components |  |  |
|  |  |  | 2 | 5 | 10 |
| odds ratio $=1.5$ |  |  |  |  |  |
| Data-set 1(FLAG) |  |  |  |  |  |
| Model 4 | 100 | $70.1{ }^{\text {b }}$ | $99.5{ }^{\text {b }}$ | $99.5{ }^{\text {b }}$ | $99.3{ }^{\text {b }}$ |
| ( "Western" pattern | 300 | 100 | 100 | 100 | 100 |
| intake positively | 600 | 100 | 100 | 100 | 100 |
| associated with | 1200 | 100 | 100 | 100 | 100 |
| disease) | 2400 | 100 | 100 | 100 | 100 |
|  | 4800 | 100 | 100 | 100 | 100 |
| Model 5 | 100 | $60.7{ }^{\text {b }}$ | $97.5{ }^{\text {b }}$ | $96.5{ }^{\text {b }}$ | $94.2{ }^{\text {b }}$ |
| ( "Western" pattern | 300 | 100 | 100 | 100 | 100 |
| intake negatively | 600 | 100 | 100 | 100 | 100 |
| associated with disease) | 1200 | 100 | 100 | 100 | 100 |
|  | 2400 | 100 | 100 | 100 | 100 |
|  | 4800 | 100 | 100 | 100 | 100 |
| Data-set 2 (ECRHS II) |  |  |  |  |  |
| Model 4 | 100 | $62.4{ }^{\text {b }}$ | $96.2^{\text {b }}$ | $97.1{ }^{\text {b }}$ | $93.9{ }^{\text {b }}$ |
| ( "Western" pattern | 300 | 100 | 100 | 100 | 100 |
| intake positively | 600 | 100 | 100 | 100 | 100 |
| associated with disease) | 1200 | 100 | 100 | 100 | 100 |
|  | 2400 | 100 | 100 | 100 | 100 |
|  | 4800 | 100 | 100 | 100 | 100 |
| Model 5 | 100 | $57.7{ }^{\text {b }}$ | $91.7^{\text {b }}$ | $91.5{ }^{\text {b }}$ | $86.1{ }^{\text {b }}$ |
| ( "Western" pattern | 300 | 100 | 100 | 100 | 100 |
| intake negatively | 600 | 100 | 100 | 100 | 100 |
| associated with disease) | 1200 | 100 | 100 | 100 | 100 |
|  | 2400 | 100 | 100 | 100 | 100 |
|  | 4800 | 100 | 100 | 100 | 100 |
| odds ratio=1.3 |  |  |  |  |  |
| Data-set 1(FLAG) |  |  |  |  |  |
| Model 4 | 100 | $44.0{ }^{\text {b }}$ | $97.1{ }^{\text {b }}$ | $97.8{ }^{\text {b }}$ | $94.2{ }^{\text {b }}$ |
| ( "Western" pattern | 300 | 100 | 100 | 100 | 100 |
| intake positively | 600 | 100 | 100 | 100 | 100 |
| associated with | 1200 | 100 | 100 | 100 | 100 |
| disease) | 2400 | 100 | 100 | 100 | 100 |
|  | 4800 | 100 | 100 | 100 | 100 |
| Model 5 | 100 | $23.8{ }^{\text {b }}$ | $92.8{ }^{\text {b }}$ | $92.9{ }^{\text {b }}$ | $84.8{ }^{\text {b }}$ |
| ( "Western" pattern | 300 | 99.9 | 100 | 100 | 100 |
| intake negatively | 600 | 100 | 100 | 100 | 100 |
| associated with disease) | 1200 | 100 | 100 | 100 | 100 |
|  | 2400 | 100 | 100 | 100 | 100 |
|  | 4800 | 100 | 100 | 100 | 100 |


| Data-set 2 (ECRHS II) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Model 4 | 100 | $55.5{ }^{\text {b }}$ | $77.6{ }^{\text {b }}$ | $74.3{ }^{\text {b }}$ | $66.6{ }^{\text {b }}$ |
| ("Western" pattern | 300 | 100 | 99.9 | 100 | 99.9 |
| intake positively | 600 | 100 | 100 | 100 | 100 |
| associated with | 1200 | 100 | 100 | 100 | 100 |
| disease) | 2400 | 100 | 100 | 100 | 100 |
|  | 4800 | 100 | 100 | 100 | 100 |
| Model 5 | 100 | $34.1{ }^{\text {b }}$ | $57.5{ }^{\text {b }}$ | $55.3{ }^{\text {b }}$ | $55.5{ }^{\text {b }}$ |
| ( "Western" pattern | 300 | 100 | 100 | 100 | 100 |
| intake negatively | 600 | 100 | 100 | 100 | 100 |
| associated with | 1200 | 100 | 100 | 100 | 100 |
| disease) | 2400 | 100 | 100 | 100 | 100 |
|  | 4800 | 100 | 100 | 100 | 100 |
| odds ratio=1.1 |  |  |  |  |  |
| Data-set 1(FLAG) |  |  |  |  |  |
| Model 4 | 100 | 1.5 | $40.1{ }^{\text {b }}$ | $28.2{ }^{\text {b }}$ | $18.0{ }^{\text {b }}$ |
| ( "Western" pattern | 300 | 49.2 | $95.9{ }^{\text {b }}$ | $94.9{ }^{\text {b }}$ | 91.0 |
| intake positively | 600 | 95.6 | 99.9 | 100 | 100 |
| associated with | 1200 | 99.9 | 100 | 100 | 100 |
| disease) | 2400 | 100 | 100 | 100 | 100 |
|  | 4800 | 100 | 100 | 100 | 100 |
| Model 5 | 100 | $1.5{ }^{\text {b }}$ | $39.2{ }^{\text {b }}$ | $29.9{ }^{\text {b }}$ | $29.8{ }^{\text {b }}$ |
| ("Western" pattern | 300 | 30.1 | 91.9 | 92.3 | 85.0 |
| intake negatively | 600 | 84.0 | 100 | 100 | 99.9 |
| associated with disease) | 1200 | 100 | 100 | 100 | 100 |
|  | 2400 | 100 | 100 | 100 | 100 |
|  | 4800 | 100 | 100 | 100 | 100 |
| Data-set 2 (ECRHS II) |  |  |  |  |  |
| Model 4 | 100 | $1.8{ }^{\text {b }}$ | $23.9{ }^{\text {b }}$ | $16.7{ }^{\text {b }}$ | $9.8{ }^{\text {b }}$ |
| ( "Western" pattern | 300 | $30.9{ }^{\text {b }}$ | $69.7{ }^{\text {b }}$ | $62.1{ }^{\text {b }}$ | $51.6{ }^{\text {b }}$ |
| intake positively | 600 | 74.9 | $94.0{ }^{\text {b }}$ | $92.2{ }^{\text {b }}$ | $88.0{ }^{\text {b }}$ |
| associated with disease) | 1200 | 99.0 | 99.9 | 99.9 | 99.9 |
|  | 2400 | 100 | 100 | 100 | 100 |
|  | 4800 | 100 | 100 | 100 | 100 |
| Model 5 | 100 | $2.1{ }^{\text {b }}$ | $19.8{ }^{\text {b }}$ | $16.2^{\text {b }}$ | $8.7{ }^{\text {b }}$ |
| ( "Western" pattern | 300 | 25.1 | 64.3 | 66.1 | 57.0 |
| intake negatively | 600 | 71.7 | 90.7 | 93.9 | 99.6 |
| associated with | 1200 | 99.3 | 99.8 | 100 | 99.9 |
| disease) | 2400 | 100 | 100 | 100 | 100 |
|  | 4800 | 100 | 100 | 100 | 100 |

${ }^{\text {a }}$ Data-set 1 is from the FLAG survey (Shaheen et al., 2001); Data-set 2 is from the UK ECRHS II survey (Hooper et al., 2010). In Model 4, foods being included in a "Western" pattern (30 in Data-set 1 and 10 in Data-set 2 - see paragraph 4.9.2 and 4.9.3) are used in each replication, with pattern intake being positively associated with disease. In Model 5, foods being included in a "Western" pattern (30 in Data-set 1 and 10 in Data-set 2 - see paragraph 4.9.2 and 4.9.3) are used in each replication, with pattern intake being negatively associated.

### 5.3 Power and FDR with which ESFA and PCA could identify specific combinations of foods between diet and disease

Our multiple null hypotheses are evaluated in tables 5.3.1, 5.3.2, and 5.3 .3 when 1 in 7 or 1 in 20 foods or a simplified "Western" dietary pattern were causally linked with disease.

In each scenario, when food intakes were causally associated with disease, ESFA had lower false discovery rate than PCA for identifying the specific combination of foods (Table 5.3.1 and 5.3.2). EFSA had greater power in all simulations except in some (17/242) of the simulations where the sample size was low (300 and 600).

When a simplified "Western" pattern intake was causally linked with disease, ESFA also had a higher power than PCA for identifying specific combination of foods (Table 5.3.3), but PCA had a lower false discovery rate (with also lower power).

Simulation results for average percentages of power and FDR estimates of ESFA didn't present a symmetrical pattern between model 1 (all selected food intakes are negatively associated with disease) and model 3 (all selected food intakes are positively associated with disease), and we get intermediate power estimates for Model 2 (half the selected food intakes are negatively associated and half are positively associated with disease). Simulation results for the ESFA method which were presented in Table 5.3 .3 were more symmetrical compared to Tables 5.3.1 and 5.3.2.

Power and FDR of PCA for detecting specific effects of foods on disease is roughly the same or higher when all effects of foods are positive compared to when all effects are negative. In addition power and FDR is less when there is a mixture of positive and negative effects compared to when all effects of foods on disease are positive. Lower or roughly the same power is observed when all effect of foods are negative compared to where there is a mixture of positive and negative effects of foods on disease. In addition, power and FDR increases as the number of principal components and the size of the sample increases. Simulation results for PCA method which were presented in Table 5.3.3 presented an almost symmetrical pattern.

Table 5.3.1. $\quad$ Power $^{c}$ and false discovery rate (FDR) ${ }^{\text {d }}$ estimates (\%) of exhaustive single food analysis (ESFA) and principal components analysis (PCA) for detecting the foods that are causally linked to disease (all estimates of power and FDR have standard error $<0.5 \%$ ). In models 1-3, one in seven foods ( 30 in data-set 1 and 10 in data-set 2) are selected at random in each replication from the foods on the FFQ.

| Data-set and model ${ }^{\text {a }}$ | Sample Size | ESFA |  | PCA <br> Number of Components |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Power | FDR | 2 |  | 5 |  | 10 |  |
|  |  |  |  | Power | FDR | Power | FDR | Power | FDR |
| Data-set 1 (FLAG) |  |  |  |  |  |  |  |  |  |
| Model 1 <br> (All foods negatively associated with disease) | 300 | 13.1 | 28.9 | 35.5 | 85.7 | 46.1 | 85.3 | $55.2{ }^{\text {b }}$ | 85.5 |
|  | 600 | 42.9 | 50.1 | 34.3 | 86.3 | 47.1 | 86.9 | 56.8 | 86.1 |
|  | 1200 | 72.6 | 64.3 | 36.2 | 86.1 | 49.5 | 86.1 | 58.6 | 85.8 |
|  | 2400 | 87.6 | 73.2 | 37.2 | 86.2 | 49.6 | 86.3 | 59.3 | 86.1 |
|  | 4800 | 93.9 | 78.9 | 38.2 | 86.5 | 51.5 | 86.1 | 60.9 | 86.0 |
| Model 2 <br> (Half foods positively, half negatively associated with disease) | 300 | 49.1 | 70.3 | $14.6{ }^{\text {b }}$ | 42.3 | $20.7{ }^{\text {b }}$ | 63.9 | $24.6{ }^{\text {b }}$ | 76.1 |
|  | 600 | 76.9 | 77.4 | 19.4 | 54.1 | 26.9 | 73.8 | 33.5 | 82.6 |
|  | 1200 | 89.7 | 80.3 | 26.0 | 67.7 | 34.2 | 82.9 | 40.2 | 85.2 |
|  | $2400$ | 95.0 | 82.2 | 29.7 | 75.0 | 38.6 | 84.3 | 46.3 | 85.8 |
|  | 4800 | 97.3 | 83.6 | 33.5 | 80.3 | 41.7 | 86.1 | 51.2 | 85.9 |
| Model 3 <br> (All foods positively associated with disease) | 300 | 55.3 | 71.8 | 35.0 | 85.2 | 47.8 | 85.8 | 55.2 | 85.4 |
|  | 600 | 77.5 | 77.8 | 36.0 | 86.4 | 48.0 | 86.2 | 57.8 | 85.7 |
|  | 1200 | 88.6 | 80.5 | 36.8 | 86.2 | 49.6 | 86.2 | 59.0 | 85.9 |
|  | 2400 | 93.8 | 82.6 | 37.4 | 86.5 | 50.0 | 86.3 | 60.2 | 85.9 |
|  | 4800 | 98.7 | 83.5 | 39.0 | 86.4 | 50.1 | 86.4 | 60.2 | 86.1 |
| Data-set 2 (ECRHS II) |  |  |  |  |  |  |  |  |  |
| Model 1 <br> (All foods negatively associated with disease) | 300 | 21.7 | 37.8 | 21.0 | 46.5 | $29.3{ }^{\text {b }}$ | 64.9 | $34.4{ }^{\text {b }}$ | 75.4 |
|  | 600 | 62.7 | 53.9 | 29.2 | 62.9 | 39.6 | 76.3 | 48.9 | 82.2 |
|  | 1200 | 90.1 | 66.1 | 35.8 | 73.6 | 49.1 | 83.1 | 60.5 | 84.1 |
|  | $2400$ | $97.9$ | 74.7 | $41.1$ | 79.9 | $56.1$ | 85.0 | 69.1 | 85.0 |
|  | 4800 | 99.5 | 79.6 | 45.3 | 83.1 | 61.4 | 85.7 | 75.2 | 85.4 |
| Model 2 <br> (Half foods positively, half negatively associated with disease) | 300 | 35.8 | 38.1 | 17.5 | 42.9 | 27.1 | 66.4 | 35.4 | 77.9 |
|  | 600 | 70.6 | 54.9 | 24.6 | 57.0 | 39.6 | 77.4 | 48.5 | 83.3 |
|  | 1200 | 90.1 | 67.2 | 31.7 | 68.1 | 45.6 | 83.3 | 60.0 | 84.5 |
|  | 2400 | 96.2 | 75.5 | 38.3 | 75.9 | 52.6 | 85.3 | 68.3 | 85.3 |
|  | 4800 | 98.2 | 80.0 | 42.9 | 80.6 | 58.7 | 85.9 | 74.0 | 85.6 |
| Model 3 <br> (All foods positively associated with disease) | 300 | 53.3 | 48.0 | 27.2 | 58.6 | 39.5 | 75.3 | 47.9 | 80.3 |
|  | 600 | 81.1 | 63.6 | 33.9 | 70.9 | 47.6 | 81.7 | 59.2 | 83.4 |
|  | 1200 | 92.9 | 73.2 | 39.5 | 78.3 | 54.4 | 84.5 | 68.0 | 84.6 |
|  | 2400 | 96.9 | 78.8 | 43.8 | 82.6 | 59.5 | 85.5 | 74.1 | 85.2 |
|  | 4800 | 98.4 | 81.7 | 46.8 | 84.3 | 63.9 | 85.9 | 78.2 | 85.9 |

[^1]Table 5.3.2. $\quad$ Power $^{c}$ and false discovery rate (FDR) ${ }^{\text {d }}$ estimates (\%) of exhaustive single food analysis (ESFA) and principal components analysis (PCA) for detecting the foods that are causally linked to disease (all estimates of power and FDR have standard error $<0.5 \%$ ). In models $1-3$, one in twenty foods ( 10 in data-set 1 and 4 in data-set 2 ) are selected at random in each replication from the foods on the FFQ.

| Data-set and model $^{\text {a }}$ | Sample <br> Size | ESFA |  | PCA <br> Number of Components |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | 2 |  | 5 |  | 10 |  |
|  |  | Power | FDR | Power | FDR | Power | FDR | Power | FDR |
| Data-set 1 (FLAG) |  |  |  |  |  |  |  |  |  |
| Model 1 <br> (All foods negatively associated with disease) | 300 | 6.5 | 22.7 | $21.5{ }^{\text {b }}$ | 60.6 | $26.5{ }^{\text {b }}$ | 70.8 | $29.9{ }^{\text {b }}$ | 80.5 |
|  | 600 | 38.2 | 52.1 | 27.3 | 77.2 | 33.1 | 84.3 | 38.0 | 89.5 |
|  | 1200 | 80.1 | 72.8 | 31.3 | 88.0 | 39.4 | 91.6 | 46.6 | 93.6 |
|  | 2400 | 95.7 | 84.8 | 33.7 | 92.9 | 44.0 | 94.0 | 50.9 | 94.7 |
|  | 4800 | 99.1 | 90.2 | 36.1 | 94.5 | 46.2 | 95.0 | 54.7 | 94.9 |
| Model 2 <br> (Half foods positively, half negatively associated with disease) | 300 | 12.1 | 21.8 | 10.8 | 34.4 | $14.7{ }^{\text {b }}$ | 52.5 | $18.6{ }^{\text {b }}$ | 70.0 |
|  | 600 | 40.4 | 41.5 | 15.7 | $48 . .9$ | 21.1 | 69.9 | 26.7 | 84.9 |
|  | 1200 | 73.6 | 63.1 | 21.5 | 64.6 | 28.0 | 82.9 | 35.5 | 91.8 |
|  | - 2400 | 91.1 | 79.9 | 26.1 | 75.4 | 34.7 | 90.5 | 42.6 | 94.5 |
|  | 4800 | 96.6 | 88.3 | 30.6 | 85.0 | 39.8 | 93.6 | 49.0 | 94.8 |
| Model 3 <br> (All foods positively associated with disease) | 300 | 41.2 | 50.3 | 26.6 | 71.6 | 33.7 | 81.0 | 39.5 | 87.2 |
|  | 600 | 76.4 | 68.9 | 30.2 | 83.4 | 39.4 | 90.3 | 46.9. | 92.7 |
|  | 1200 | 94.3 | 82.2 | 33.0 | 95.1 | 43.6 | 93.7 | 52.1 | 94.2 |
|  | 2400 | 98.5 | 89.1 | 35.3 | 93.3 | 45.6 | 94.9 | 55.9 | 94.8 |
|  | 4800 | 99.5 | 92.7 | 38.0 | 94.4 | 48.0 | 95.2 | 58.0 | 95.0 |
| Data-set 2 (ECRHS II) |  |  |  |  |  |  |  |  |  |
| Model 1 <br> (All foods negatively associated with disease) | 300 | 1.2 | 8.5 | $8.8{ }^{\text {b }}$ | 24.1 | $12.2{ }^{\text {b }}$ | 37.5 | $12.7{ }^{\text {b }}$ | 51.0 |
|  | 600 | 11.2 | 19.9 | $13.8{ }^{\text {b }}$ | 36.6 | $17.9{ }^{\text {b }}$ | 53.3 | $22.7{ }^{\text {b }}$ | 68.6 |
|  | 1200 | 50.7 | 38.6 | 20.9 | 54.1 | 26.5 | 71.6 | 31.9 | 83.3 |
|  | 2400 | 85.9 | 62.0 | 27.5 | 74.1 | 34.0 | 85.5 | 41.7 | 93.1 |
|  | 4800 | 96.2 | 81.6 | 33.1 | 83.5 | 40.6 | 92.1 | 49.8 | 96.5 |
| Model 2 <br> (Half foods positively, half negatively associated with disease) | 300 | 29.7 | 27.8 | 12.8 | 32.2 | 21.2 | 55.2 | 29.2 | 72.2 |
|  | 600 | 64.7 | 46.3 | 19.2 | 46.3 | 30.7 | 71.6 | 43.5 | 84.9 |
|  | 1200 | 90.3 | 65.1 | 26.3 | 60.3 | 40.7 | 84.1 | 56.8 | 90.5 |
|  | - 2400 | 97.9 | 80.4 | 34.0 | 74.2 | 49.4 | 90.7 | 67.2 | 92.4 |
|  | 4800 | 99.5 | 87.9 | 39.4 | 83.2 | 56.9 | 93.1 | 74.2 | 93.3 |
| Model 3 <br> (All foods positively associated with disease) | 300 | 51.5 | 35.5 | 20.5 | 43.3 | 31.3 | 65.2 | 41.7 | 79.1 |
|  | 600 | 86.6 | 56.1 | 28.0 | 58.6 | 40.6 | 80.2 | 54.9 | 88.5 |
|  | 1200 | 98.3 | 74.7 | 34.6 | 71.9 | 49.4 | 88.8 | 66.0 | 91.7 |
|  | 2400 | 99.8 | 85.7 | 39.9 | 82.3 | 56.1 | 92.3 | 73.9 | 92.8 |
|  | 4800 | 99.9 | 90.2 | 44.6 | 88.2 | 61.3 | 93.6 | 78.2 | 93.5 |

[^2]Table 5.3.3. $\quad$ Power $^{\mathrm{c}}$ and false discovery rate (FDR) ${ }^{\text {d }}$ estimates (\%) of exhaustive single food analysis (ESFA) principal components analysis (PCA) for detecting the foods that deemed to constitute a simplified "Western" dietary pattern and are causally linked to disease (all estimates of power and fdr have standard error $<\mathbf{0 . 5 \%}$ ). In model 4, foods being included in a simplified "Western" pattern are used in each replication.

| Data-set and model ${ }^{\text {a }}$ | Sample <br> Size | ESFA |  | PCANumber of Components |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Power | FDR | 2 |  | 5 |  | 10 |  |
|  |  |  |  | Power | FDR | Power | FDR | Power | FDR |
| Data-set 1 (FLAG) |  |  |  |  |  |  |  |  |  |
| Model 4 | 300 | 88.6 | 76.7 | 55.3 | 82.4 | 74.8 | 80.7 | 79.8 | 81.7 |
| ("Western" | 600 | 97.6 | 80.5 | 67.5 | $77.9^{\text {c }}$ | 81.6 | $79.2^{\text {c }}$ | 83.3 | 81.3 |
| pattern intake | 1200 | 99.7 | 82.8 | 78.2 | $74.3{ }^{\text {c }}$ | 86.9 | $77.4{ }^{\text {c }}$ | 86.4 | $80.7{ }^{\text {c }}$ |
| positively associated with | 2400 | 99.1 | 84.1 | 84.5 | $72.2{ }^{\text {c }}$ | 90.9 | $76.1{ }^{\text {c }}$ | 87.7 | $80.4{ }^{\text {c }}$ |
| disease) | 4800 | 100 | 84.9 | 87.3 | $71.5{ }^{\text {c }}$ | 93.1 | $75.2{ }^{\text {c }}$ | 87.7 | $80.4{ }^{\text {c }}$ |
| Model 5 | 300 | 92.2 | 77.7 | 54.0 | 82.6 | 75.6 | 81.4 | 80.8 | 82.3 |
| ("Western" | 600 | 98.7 | 81.1 | 67.9 | $77.8^{\text {c }}$ | 82.2 | $79.4{ }^{\text {c }}$ | 84.2 | 81.6 |
| pattern intake | 1200 | 99.9 | 83.3 | 78.4 | $74.2{ }^{\text {c }}$ | 87.1 | $77.5{ }^{\text {c }}$ | 86.5 | $80.9{ }^{\text {c }}$ |
| negatively associated | 2400 | 100 | 84.3 | 84.4 | $72.3{ }^{\text {c }}$ | 90.7 | $76.0{ }^{\text {c }}$ | 87.8 | $80.5{ }^{\text {c }}$ |
| with disease) | 4800 | 100 | 84.9 | 87.4 | $71.5{ }^{\text {c }}$ | 93.2 | $75.2{ }^{\text {c }}$ | 87.9 | $80.5{ }^{\text {c }}$ |
| Data-set 2 (ECRHS II) |  |  |  |  |  |  |  |  |  |
| Model 4 | 300 | 96.1 | 56.3 | 60.8 | 74.7 | 73.4 | 72.0 | 84.5 | 73.3 |
| ( "Western" | 600 | 99.9 | 67.1 | 63.6 | 74.3 | 76.6 | 72.9 | 89.5 | 75.3 |
| pattern intake | 1200 | 100 | 73.0 | 66.5 | 74.2 | 79.9 | 73.1 | 92.6 | 76.4 |
| positively <br> associated with | 2400 | 100 | 76.0 | 68.6 | $74.0{ }^{\text {c }}$ | 84.5 | $72.6{ }^{\text {c }}$ | 93.6 | 77.0 |
| disease) | 4800 | 100 | 77.3 | 70.3 | $73.7^{\text {c }}$ | 89.8 | $71.6{ }^{\text {c }}$ | 94.2 | 77.3 |
| Model 5 <br> ("Western" pattern intake negatively associated with disease) | 300 | 96.8 | 60.3 | 61.0 | 74.6 | 74.1 | 74.0 | 84.2 | 74.0 |
|  | 600 | 99.9 | 68.2 | 63.9 | 74.4 | 77.1 | 74.8 | 89.2 | 75.8 |
|  | 1200 | 100 | 73.3 | 66.4 | 74.2 | 80.5 | 74.5 | 92.2 | 76.7 |
|  | 2400 | 100 | 76.0 | 68.7 | $73.9{ }^{\text {c }}$ | 85.7 | $73.3{ }^{\text {c }}$ | 94.1 | 77.1 |
|  | 4800 | 100 | 77.4 | 70.3 | $73.7^{\text {c }}$ | 91.2 | $71.9{ }^{\text {c }}$ | 94.8 | 77.3 |

[^3]
### 5.4 Power and FDR with which ESFA and PCA could identify specific combinations of foods between diet and disease adjusted in three different ways

Table 5.4.1, 5.4.2, and 5.4.3 presents average estimates of percentages of power and false discovery rates for different sample sizes when exhaustive single food analysis is employed for 3 different ways of adjustment (see paragraphs 4.3.2, 4.3.3 and 4.3.4).

Attempting to control the FDR of ESFA by adjusting for principal components of diet or propensity scores was not successful for either datasets, especially for large sample sizes. However, adjusting for other foods that were significant in an unadjusted analysis controlled the FDR at around the nominal $20 \%$ level, though with some loss of power, particularly with low sample sizes.

Simulation results for average percentages of power estimates of adjusted ESFA didn't present a symmetrical pattern between model 1 (all selected food intakes are negatively associated with disease) and model 3 (all selected food intakes are positively associated with disease), and we get intermediate power estimates for Model 2 (half the selected food intakes are negatively associated and half are positively associated with disease) for small sample size scenarios (300, 600 and 1200). However for larger sample size scenarios (2400, 4800) they are roughly the same. Average parentages of FDR for the adjusted ESFA, exceeded the nominal level of $20 \%$ for the F.L.A.G survey data-set for specific sample size scenarios (300, 600,1200 ). However, as the sample size increases above 2400 , FDR was controlled around $20 \%$. Not any other specific patterns across Tables 5.4.1, 5.4.2 and 5.4.3 were observed.

Simulation results of power and FDR for ESFA adjusted for 5 principal components were qualitatively similar with the simulation results that were observer for the unadjusted ESFA. (Tables in paragraphs 5.3). No clear and consistent simulation patterns for the average percentages of power and FDR were observed when the ESFA method was adjusted for propensity scores.

Table 5.4.1. $\quad$ Power $^{c}$ and false discovery rate (FDR) ${ }^{\text {d }}$ estimates (\%) of exhaustive single food analysis (ESFA) with different methods of adjustment for other foods (all estimates of power and FDR have standard error $<\mathbf{0 . 5 \%}$ ). In models $\mathbf{1 - 3}$, one in seven foods ( $\mathbf{3 0}$ in data-set 1 and 10 in data-set 2) are selected at random in each replication from the foods on the FFQ. FDR of Simes procedure at $\mathbf{2 0 \%}$.

| Data-set and model ${ }^{\text {a }}$ | Sample <br> Size | Adjusted for 5 principal components |  | Adjusted for foods that are significant in unadjusted analysis |  | Adjusted for propensity scores |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Power | FDR | Power | FDR | Power | FDR |
| Data-set 1 (FLAG) |  |  |  |  |  |  |  |
| Model 1 <br> (All foods negatively associated with disease) | 300 | 1.7 | 27.5 | 4.8 | 29.4 | 0.9 | 8.5 |
|  | 600 | 19.4 | 44.5 | 24.5 | 48.8 | 17.5 | 33.6 |
|  | 1200 | 59.6 | 53.7 | 50.9 | 35.3 | 66.3 | 47.1 |
|  | 2400 | 85.5 | 64.2 | 78.2 | 27.2 | 99.2 | 58.5 |
|  | 4800 | 95.1 | 72.6 | 91.4 | 24.3 | 98.6 | 69.7 |
| Model 2 <br> (Half foods positively, half negatively associated with disease) | 300 | 6.5 | 19.1 | 5.0 | 18.5 | 0.1 | 1.2 |
|  | 600 | 35.0 | 35.8 | 30.5 | 33.4 | 2.3 | 3.3 |
|  | 1200 | 71.5 | 51.5 | 67.5 | 31.8 | 35.5 | 8.0 |
|  | 2400 | 89.9 | 65.0 | 88.7 | 27.8 | 79.2 | 14.5 |
|  | 4800 | 96.4 | 73.9 | 96.1 | 25.7 | 93.9 | 24.8 |
| Model 3 <br> (All foods positively associated with disease) | 300 | 7.3 | 26.8 | 5.2 | 17.2 | 0.2 | 0.8 |
|  | 600 | 33.0 | 43.4 | 24.6 | 33.7 | 1.8 | 11.2 |
|  | 1200 | 67.5 | 57.3 | 64.5 | 41.9 | 24.9 | 19.4 |
|  | 2400 | 87.9 | 68.5 | 86.6 | 30.4 | 91.2 | 54.0 |
|  | 4800 | 95.6 | 75.6 | 95.3 | 27.4 | 94.2 | 54.9 |
| Data-set 2 (ECRHS II) |  |  |  |  |  |  |  |
| Model 1 <br> (All foods negatively associated with disease) | 300 | 9.6 | 30.2 | 4.2 | 15.9 | 5.7 | 7.7 |
|  | 600 | 40.1 | 44.1 | 22.7 | 19.8 | 28.4 | 10.8 |
|  | 1200 | 77.5 | 55.4 | 67.5 | 19.7 | 73.6 | 16.0 |
|  | 2400 | 94.3 | 66.8 | 93.3 | 19.0 | 96.4 | 22.8 |
|  | 4800 | 98.6 | 74.8 | 99.2 | 18.1 | 99.7 | 35.5 |
| Model 2 <br> (Half foods positively, half negatively associated with disease) | 300 | 25.4 | 29.9 | 15.4 | 15.2 | 4.1 | 6.4 |
|  | 600 | 61.5 | 45.7 | 47.2 | 19.6 | 26.5 | 10.3 |
|  | 1200 | 86.9 | 59.9 | 82.1 | 20.0 | 72.2 | 16.1 |
|  | 2400 | 96.0 | 70.8 | 96.3 | 19.9 | 94.5 | 23.7 |
|  | 4800 | 98.7 | 77.7 | 99.0 | 20.4 | 99.1 | 33.7 |
| Model 3 <br> (All foods negatively associated with disease) | 300 | 32.3 | 35.1 | 23.0 | 16.8 | 3.0 | 5.5 |
|  | 600 | 67.0 | 52.4 | 60.7 | 20.7 | 29.5 | 11.7 |
|  | 1200 | 88.0 | 66.0 | 90.7 | 21.4 | 66.8 | 22.5 |
|  | 2400 | 95.7 | 74.7 | 98.8 | 21.4 | 93.2 | 37.7 |
|  | 4800 | 98.2 | 79.5 | 99.9 | 21.4 | 98.6 | 58.0 |

[^4]Table 5.4.2. $\quad$ Power $^{c}$ and false discovery rate (FDR) ${ }^{\text {d }}$ estimates (\%) of exhaustive single food analysis (ESFA) with different methods of adjustment for other foods (all estimates of power and FDR have standard error $<\mathbf{0 . 5 \%}$ ). In models $\mathbf{1 - 3}$, one in twenty foods ( 10 in data-set 1 and 4 in data-set 2) are selected at random in each replication from the foods on the FFQ. FDR of Simes procedure at $\mathbf{2 0 \%}$.

| Data-set and model ${ }^{\text {a }}$ | Sample <br> Size | Adjusted for 5 principal components |  | Adjusted for foods that are significant in unadjusted analysis |  | Adjusted for propensity scores |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Power | FDR | Power | FDR | Power | FDR |
| Data-set 1 (FLAG) |  |  |  |  |  |  |  |
| Model 1 <br> (All foods negatively associated with disease) | 300 | 0.7 | 13.4 | 1.1 | 11.2 | 0.7 | 10.2 |
|  | 600 | 8.7 | 26.7 | 5.9 | 21.5 | 4.6 | 13.8 |
|  | 1200 | 46.8 | 37.7 | 31.0 | 27.4 | 38.1 | 24.5 |
|  | 2400 | 82.9 | 54.9 | 70.3 | 22.8 | 83.3 | 36.4 |
|  | 4800 | 95.4 | 73.0 | 89.1 | 21.3 | 96.9 | 53.0 |
| Model 2 <br> (Half foods positively, half negatively associated with disease) | 300 | 7.7 | 13.6 | 6.0 | 10.5 | 0.1 | 5.8 |
|  | 600 | 32.8 | 26.8 | 27.9 | 21.1 | 4.1 | 5.4 |
|  | 1200 | 69.3 | 45.6 | 63.6 | 23.0 | 35.4 | 10.4 |
|  | 2400 | 91.1 | 67.4 | 86.0 | 23.2 | 77.5 | 14.1 |
|  | 4800 | 97.4 | 82.8 | 94.6 | 20.5 | 93.4 | 21.0 |
| Model 3 <br> (All foods positively associated with disease) | 300 | 12.7 | 16.1 | 11.4 | 19.3 | 0.1 | 2.3 |
|  | 600 | 48.7 | 31.5 | 44.3 | 29.5 | 3.7 | 4.3 |
|  | 1200 | 83.6 | 52.6 | 79.3 | 28.1 | 41.2 | 9.5 |
|  | 2400 | 96.5 | 74.5 | 93.4 | 24.4 | 87.1 | 14.7 |
|  | 4800 | 99.2 | 86.4 | 97.5 | 21.5 | 98.1 | 25.6 |

Data-set 2 (ECRHS II)

| Model 1 | 300 | 9.6 | 30.2 | 4.2 | 15.9 | 5.7 | 7.7 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| (All foods | 600 | 40.1 | 44.1 | 22.7 | 19.8 | 28.4 | 10.8 |
| negatively | 1200 | 77.5 | 55.4 | 67.5 | 19.7 | 73.6 | 16.0 |
| associated | with disease) | 2400 | 94.3 | 66.8 | 93.3 | 19.0 | 96.4 |
|  | 4800 | 98.6 | 74.8 | 99.2 | 18.1 | 99.7 | 35.5 |
| Model 2 | 300 | 25.4 | 29.9 | 15.4 | 15.2 | 4.1 | 6.4 |
| (Half foods | 600 | 61.5 | 45.7 | 47.2 | 19.6 | 26.5 | 10.3 |
| positively, half | 1200 | 86.9 | 59.9 | 82.1 | 20.0 | 72.2 | 16.1 |
| negatively | 2400 | 96.0 | 70.8 | 96.3 | 19.9 | 94.5 | 23.7 |
| associated | with disease) | 4800 | 98.7 | 77.7 | 99.0 | 20.4 | 99.1 |
| Model 3 | 300 | 32.3 | 35.1 | 23.0 | 16.8 | 3.0 | 33.7 |
| (All foods | 600 | 67.0 | 52.4 | 60.7 | 20.7 | 29.5 | 11.7 |
| positively | 1200 | 88.0 | 66.0 | 90.7 | 21.4 | 66.8 | 22.5 |
| associated | with disease) | 2400 | 95.7 | 74.7 | 98.8 | 21.4 | 93.2 |
|  | 4800 | 98.2 | 79.5 | 99.9 | 21.4 | 98.6 | 58.7 |

[^5]Table 5.4.3. $\quad$ Power $^{c}$ and false discovery rate (FDR) ${ }^{\text {d }}$ estimates (\%) of exhaustive single food analysis (ESFA) with different methods of adjustment for other foods (all estimates of power and FDR have standard error $<0.5 \%$ ). In model 4, foods being included in a "Western" pattern are used in each replication. FDR of Simes procedure at $\mathbf{2 0 \%}$.

| Data-set and model $^{\text {a }}$ | Sample Size | Adjusted for 5 principal components |  | Adjusted for foods that are significant in unadjusted analysis |  | Adjusted for propensity scores |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Power | FDR | Power | FDR | Power | FDR |
| Data-set 1(FLAG) |  |  |  |  |  |  |  |
| Model 4 | 300 | 4.3 | 52.8 | 2.0 | 18.5 | 0.5 | 21.3 |
| ("Western" | 600 | 13.5 | 70.1 | 32.5 | 33.4 | 2.1 | 79.5 |
| pattern intake | 1200 | 32.7 | 75.6 | 57.5 | 38.8 | 19.7 | 77.7 |
| positively <br> associated with | 2400 | 56.1 | 78.0 | 72.7 | 35.1 | 54.2 | 74.3 |
| disease) | 4800 | 72.5 | 80.6 | 94.4 | 22.9 | 76.4 | 76.8 |
| Model 5 | 300 | 3.2 | 43.9 | 3.2 | 22.5 | 1.1 | 13.8 |
| ("Western" | 600 | 12.1 | 65.8 | 24.1 | 26.9 | 27.8 | 34.8 |
| pattern intake | 1200 | 33.8 | 71.4 | 44.9 | 47.4 | 70.1 | 40.0 |
| negatively associated with | 2400 | 61.2 | 75.2 | 73.7 | 29.2 | 70.3 | 39.1 |
| disease) | 4800 | 79.2 | 79.0 | 94.5 | 20.8 | 76.6 | 76.8 |
| Data-set 2 (ECRHS II) |  |  |  |  |  |  |  |
| Model 4 | 300 | 15.6 | 55.1 | 12.4 | 23.9 | 3.6 | 5.2 |
| ( "Western" | 600 | 33.8 | 66.9 | 34.1 | 21.9 | 26.9 | 18.1 |
| pattern intake | 1200 | 50.6 | 70.3 | 65.6 | 18.8 | 67.5 | 33.5 |
| positively <br> associated with | 2400 | 63.1 | 73.3 | 87.1 | 17.2 | 83.1 | 50.0 |
| disease) | 4800 | 71.4 | 75.7 | 96.7 | 17.5 | 89.6 | 61.3 |
| Model 5 | 300 | 14.7 | 54.5 | 9.2 | 22.5 | 6.4 | 6.1 |
| ("Western" | 600 | 33.4 | 67.6 | 27.1 | 22.9 | 38.2 | 11.1 |
| pattern intake | 1200 | 50.6 | 70.1 | 61.7 | 19.2 | 77.4 | 21.0 |
| negatively | 2400 | 63.0 | 73.4 | 90.2 | 18.1 | 95.5 | 34.4 |
| associated with disease) | 4800 | 71.8 | 75.5 | 98.7 | 17.8 | 99.7 | 50.4 |

[^6]
## 6 Analysis of G.A ${ }^{2}$.L.E.N data: effect of diet on asthma

### 6.1 Introduction

6.2 Materials and methods
6.2.1 Study design and population
6.2.2 Definitions of respiratory outcomes
6.2.3 Dietary assessment
6.2.3.1 Food Frequency Questionnaire (FFQ)
6.2.3.2 Exclusions of dietary data
6.2.4 Statistical Analysis
6.2.4.1 Dietary patterns analysis with the use of Principal Component Analysis (PCA)
6.2.4.2 Two step Exhaustive Single Food Analysis (ESFA) for multicentre data

### 6.3 Results

6.3.1 Descriptive statistics
6.3.2 Empirically derived dietary patterns with the use of Principal Component Analysis (PCA)
6.3.2.1. For all countries combined
6.3.2.2 For multicentre data
6.3.3 Dietary patterns and respiratory outcomes
6.3.4 Exhaustive single food analysis and respiratory outcomes

### 6.1 Introduction

There is accumulating evidence from observational studies that high intakes of fruit, vegetable and oily fish have a protective effect in children and adults with asthma (Oien, Storro \& Johnsen, 2010, Hodge et al., 1996, Fitzsimon et al., 2007). Additionally, asthma has been associated with dietary antioxidants and particularly with low intakes of vitamin C, D and E and selenium. (Shaheen et al., 2001, Hodge et al., 1996, Devereux, 2010). However, trials of supplementation have been unsuccessful and have provided contradictory results (Pearson et al., 2004, Fogarty et al., 2003, Devereux \& Seaton, 2005, Shaheen et al., 2007). This may be because apparent effects in the observational studies were confounded by lifestyle factors or by other dietary components. In addition, there are a number of methodological problems in analysing the effects of single food and nutrient analysis, as described in paragraph 2.1.2.

Dietary patterns empirically derived with the use of Principal Components Analysis are provided by the literature as an alternative way of investigating associations between diet and asthma. Studies have provided only weak (Bakolis et al., 2010, Butler et al., 2006) or no statistical significant evidence (Hooper et al., 2010, Shaheen et al., 2009, Varraso et al., 2009) that any of these patterns were associated with asthma. However, associations between dietary patterns and wheeze (Takaoka \& Norback, 2008), allergic rhinitis (Bakolis et al., 2010), positive skin prick tests (Hooper et al., 2010) and COPD (Varraso et al., 2007a, Varraso et al., 2007b) have been observed. These associations are limited by conceptual and methodological disadvantages of PCA as described by the literature (Newby \& Tucker, 2004, Slattery \& Boucher, 1998) and by paragraph 2.6.

The purpose of this chapter is to provide a comprehensive analysis of the association between dietary intake in adults across Europe in relation to self-reported asthma and other respiratory and allergic symptoms. In addition, we aim to compare and interpret the results from dietary pattern analysis with the use of PCA with our two-step ESFA procedure.

### 6.2 Materials and methods

### 6.2.1 Study design and population

GA ${ }^{2}$ LEN was an EU funded network of excellence which coordinated the study of genetic and environmental risk factors for asthma in adult and adolescent population across 17 centres in 11 European countries. The GA $^{2}$ LEN follow up survey is a cross-sectional study amongst those previously contacted in baseline postal surveys that were willing to be contacted again. Invitation to follow-up was dependent on three groups of cases (those with asthma, those with sinusitis, and those with both asthma and sinusitis) and one group of controls (those with neither asthma nor sinusitis) drawn from the postal survey sample. Body and height measurements and skin prick tests (SPTs) to grass pollen, grass mix, Dermatophagoides pteronyssinus, Dermatophagoides farinae, cockroach (Blatella), olive, Alternaria, dog, Artemisia, birch, cat and Parietaria were conducted. In all centres permission to conduct this study was obtained from appropriate local ethics committees, and all participants signed a written consent form after being fully informed about the study.

### 6.2.2 Definitions of respiratory symptoms

Asthma was defined as present in those who had answered yes to having a diagnosis of asthma and either wheezing, waking up with chest tightness, waking up with shortness of breath, or waking at night with an attack of coughing in the previous 12 months. Also Chronic sinusitis (CRS) was diagnosed in those who reported that nose had been blocked, pain/pressure around forehead nose/eyes, discoloured nasal discharge (snot)/discoloured mucus in the throat, sense of smell been reduced/absent in the previous 12 months. A symptom-based definition of CRS, according to the epidemiological part of the European Position Paper on Rhinosinusitis and Nasal Polyps (EPO3S) criteria is suitable for the assessment of geographic variation in prevalence of CRS (Tomassen et al., 2011). Other respiratory symptoms included allergic rhinitis, which was defined as a positive response to the question "Do you have any nasal allergies including hay fever?" and eczema was defined as a positive response to the question "Did you ever have eczema or any kind of skin allergy?" Atopy was defined as any positive response ( 1 mm more than the diluent control) to any of the allergens tested (grass pollen, grass mix, Dermatophagoides pteronyssinus, Dermatophagoides farinae, cockroach (Blatella), olive, Alternaria, dog, Artemisia, birch, cat and Parietaria).

### 6.2.3 Dietary assessment

### 6.2.3.1 Food Frequency Questionnaire (FFQ)

The GA $^{2}$ LEN survey objective is to assess dietary intake across European countries using a single common and standardised method. It is the first study that the same standardised FFQ was administered across the European countries that took part. We designed the FFQ taking into account what other large international epidemiological studies have done to assess dietary intake in various European countries. In particular, the researchers from EPIC (Epidemiological Prospective Study in the causes of Cancer) facilitated all the FFQs used in each country. We also had valuable input from various patient associations and from lay members of the public representing each country.

FFQs recorded a consumption of 239 food items over the last 12 months ranging from never to two portions a day or more. We estimated weekly intake (g) of foods and food groups by multiplying frequency of consumption by the weight of standard portion sizes.

Prior to the FFQ being used in the GA ${ }^{2}$ LEN Nutrition survey, we successfully piloted and validated it in five countries representing different regions of Europe (Scandinavia (Helsinki), South Mediterranean (Athens), Central Europe (Brandenburg), East Europe (Lodz), North Mediterranean (Porto) (Garcia-Larsen et al., 2011). The FFQ showed a high level of repeatability for most nutrients. We also validated the FFQ comparing dietary $\mathrm{n}-3$ fatty acids against specific fatty acids in plasma. We found a good correlation between $n-3$ fatty acids in diet compared with total plasma phospholipid n-3 fatty acids and with docosahexaenoic acid. This was observed both in the entire sample (ICC 0.40) and per country. These results indicate that the GA ${ }^{2}$ LEN FFQ is an appropriate tool to estimate dietary intake for a range of nutrients across Europe regardless of cultural and linguistic differences (Garcia-Larsen et al., 2011)..

### 6.2.3.2 Exclusions of dietary data

Respondents sometimes left individual items blank on the FFQs. This was assumed to denote zero intakes of these foods unless more than $20 \%$ of items were blank, in which case the FFQ was considered incomplete, and the subject was excluded from analyses. We have included in our analysis only food items which were consumed at least 1-3 times per week from more than $2 \%$ of the individuals in our final sample (see appendix X for further details).

Based on this criterion, xilopites were excluded from our analysis. In this case we used a more a higher threshold for excluding food items than what is recommended (Willet, 1998).

### 6.2.4 Statistical Analysis

6.2.4.1 Dietary patterns analysis with the use of PCA

First, in order to derive our dietary patterns, we ran a principal component analysis for all data combined and estimated our principal component scores as we described at paragraph 2.4.

Since there is an expectation of between country heterogeneity our principal component scores tend to be correlated within countries. An alternative method which takes into account this dependence of our principal component scores is to identify our dietary patterns with the use of PCA from an overall pooled correlation matrix using a meta-analysis method proposed by Hedges and Olkin (Hedges, 1985). Specifically, in each country k for each food item $i$ and food item $j$, we evaluated the correlation matrix using the Pearson product-moment correlation coefficient $r_{k i j}$. Because the approximate distribution of $r_{k i j}$ depends strongly on the value of the population correlation $\rho_{k i j}$, each correlation coefficient was transformed using a Fisher transformation
$\left.z_{k i j}=0.5 \log \left(\left(1+r_{k i j}\right) /\left(1-r_{k i j}\right)\right)\right), i=1, . .239 \quad j=1, . .239 \quad k=1, . .10$
to give it an approximately normal distribution with asymptotic variance $1 /\left(n_{j}-3\right)$, where $n_{k}$ is the sample size for the country $k$. A weighted average of these values was then calculated

$$
\sum_{k=1}^{10} w_{k} z_{k j}=\mathrm{w}_{1} \mathrm{Z}_{1 \mathrm{ij}}+\ldots+\mathrm{w}_{10} \mathrm{z}_{10 \mathrm{ij},}, i=1, . .239 \quad j=1, . .239
$$

where the weights are $w_{k}=\frac{\left(n_{k}-3\right)}{\sum_{l=1}^{10}\left(n_{l}-3\right)}$

An inverse Fisher transformation was then applied to give a pooled correlation coefficient matrix. PCA was applied to the matrix of pooled correlation coefficients, giving us dietary pattern scores which could be used in all the 10 countries. This meta-analytic approach to PCA has previously been applied in the field of psychiatry (Smith, Mar \& Turoff, 1998, Grube, Bilder \& Goldman, 1998) and asthma epidemiology (Hooper et al., 2010).

For each country, we used multivariable logistic regression to investigate associations between the dietary patterns (in quintile groups) and respiratory outcomes adjusted for age, sex, smoking status and body mass index. We selected our confounders based on information in previous literature (Tricon et al., 2006). The effects of the dietary patterns were also adjusted for each other, because although principal components are uncorrelated, varimax rotations can introduce correlations between the dietary patterns. Regression results were pooled across countries using random effects meta-analysis, with a test for heterogeneity of regression coefficients (DerSimonian \& Laird, 1986). Heterogeneity was summarised using the $I^{2}$ statistic (Higgins et al., 2003).

All analyses were weighted to the population that took part in the postal survey. The sampling probability weights for each subject in each subset of the survey data were computed by dividing the frequency of the subject's centre and case status in the postal survey by the frequency of the same centre and case status in the subset.

### 6.2.4.2 Two step Exhaustive Single Food Analysis (ESFA) for multicentre data

In multi-centre studies, because responses tend to be correlated within centres, a method which takes into account this dependence is a multilevel model (Goldstein, 2011, Hox, 2010). Specifically, associations of self reported respiratory outcomes with each individual food in the FFQ were assessed using a random intercept logistic model. The random intercept could be thought as the combined effect of all omitted individual-specific covariates that cause some individuals to be more prone to the potential respiratory outcomes than others and is used to model the unobserved heterogeneity between countries. Random intercept models treat countries as a random representative sample from a "larger" population of countries (Rabe-Hesketh, Sophia 2006, Hox, 2010).

We employed our two-step exhaustive single food analysis framework. First we ran the analysis controlling the false discovery rate at $20 \%$. Second, we re-ran the same analysis additionally adjusting for foods that were identified as statistically significant at the first step and controlling the false discovery rate at $20 \%$. Similar analyses were conducted for associations between each individual food and asthma, sinusitis, nasal allergies, eczema and atopy. All analysis was weighted to the population that took part in the postal survey and adjusted for age, sex, smoking status and body mass index (we selected our confounders based on previous literature (Tricon et al., 2006)). Sampling probability weights were
rescaled according to a method proposed by Rabe-Hesketh in order not to induce bias in our standard estimators (Rabe-Hesketh, Sophia 2006).

Statistical analyses were conducted using STATA 12 (Stata Corporation, College Station, Texas USA).

### 6.3 Results

### 6.3.1 Descriptive statistics

Centres in Palermo (Italy), Krakow (Poland) and Skopje (Macedonia) were excluded because of the small number of cases and additionally small number of individuals who completed the FFQ questionnaire ( $\mathrm{n}=32, \mathrm{n}=0$ and $\mathrm{n}=26$ respectively). We didn't merge the German centres (Berlin, Duisberg) into one country because of socio-economic and demographic differences between their populations. Our final sample included 3057 individuals living in 15 centres in 9 countries with full data on food frequency questionnaire data and confounders. Food item entitled xilopites was excluded from our analysis due to infrequent consumption (which leads to no disease relevance) across countries; only 18 individuals (18/3057=0.4 \%) consumed xilopites (see Appendix X for further details). Further descriptive data on health outcomes and confounders are presented in Table 6.3.1.1.

Table 6.3.1.1. Description of sample

|  | Number/Total with information | Proportion (\%) unless otherwise stated |
| :---: | :---: | :---: |
| Respiratory and Allergic Outcomes |  |  |
| Asthma (\%) | 1078/3057 | 35.2 |
| Chronic Sinusitis (\%) | 595/3057 | 19.4 |
| Allergic Rhinitis (\%) | 1437/3057 | 47.0 |
| Eczema (\%) | 1760/3057 | 57.5 |
| Atopy (\%) | 1434/3057 | 46.9 |
|  |  |  |
| Confounders |  |  |
| Age (med, IQR) | 3057 | (median: 48 , IQR : 36-60) |
| Female (\%) | 1746/3057 | 57.1 |
|  |  |  |
| Smokers (\%) |  |  |
| Never | 1519/3057 | 49.6 |
| Ever | 991/3057 | 32.6 |
| Current | 539/3057 | 17.6 |
|  |  |  |
| Body Mass Index (med, IQR) | 3057 | ( median: 25.2, IQR : 22.7-28.6) |
|  |  |  |
| Country (\%) |  |  |
| Belgium(Ghent) | 146/3057 | 4.78 |


| Denmark(Odense) | $354 / 3057$ | 11.58 |
| :--- | :---: | :---: |
| Finland(Helsinki) | $155 / 3057$ | 5.07 |
| North Germany(Berlin) | $177 / 3057$ | 5.79 |
| South Germany(Duisberg) | $194 / 3057$ | 6.35 |
| Holland(Amsterdam) | $215 / 3057$ | 7.03 |
| Portugal(Coimbra) | $259 / 3057$ | 8.47 |
| Poland(Katowice, Lodz, Krakow) | $210 / 3057$ | 6.87 |
| UK(Southampton, London) | $171 / 3057$ | 5.59 |
| Sweden(Umea, Uppsala, Gothenburg) | $1176 / 3057$ | 38.47 |

### 6.3.2 Empirically derived dietary patterns with the use of PCA

### 6.3.2.1 For all countries combined

Principal component analysis was applied to the FFQ data for all countries combined. This was performed in controls only to avoid potential bias. A varimax rotation was applied to improve the interpretability of the patterns obtained. The scree plot (Figure 6.3.2.1.1) was examined to aid the choice of number of patterns, but this choice was mainly based on the principal component interpretability. We extracted the first two components (dietary patterns), which explained $14.8 \%$ of the variance in the original 238 items (food item entitled xilopites was removed due to limited observations; $0.5 \%$ of the sample ( 12 individuals in total) consumed xilopites in 4 different countries). A component score was created for each individual for each of the principal components identified. Individual foods that correlated $>0.3$ or $<-0.3$ with the varimax rotated dietary patterns (principal components) are shown in Table 6.3.2.1.1. Two patterns were derived; a "fruit, fish and vegetables" pattern (containing higher levels of fish, fruit, vegetables and boiled chicken) and a "meat, potatoes and sweets" one (containing higher levels of high fat foods, potatoes and meats).

Figure 6.3.2.1.2 presents ten different scree plots when PCA was performed for each country separately, showing different number of patterns (4 in Belgium, 5 in Denmark, 2 in Finland, 4 in North Germany, 3 in South Germany, 3 in Holland, 4 in Portugal, 2 in Poland, 4 in UK and 3 in Sweden) that could potentially be derived empirically by country. Furthermore, we calculated the mean of "fruit, fish and vegetables" and "meat, potatoes and sweets" pattern intake for all the individuals for each country separately. We observed that there is variation of the "meat, potatoes and sweets" and "fruit, fish and vegetables" mean pattern intake scores between countries (Figure 6.3.2.1.3). Therefore, using pattern scores without accounting for between country heterogeneity would lead to false estimations and conclusions.

Table 6.3.2.1.1. Foods items which correlated $>0.3$ or <-0.3 with the identified dietary patterns for all the data combined (Food items that didn't correlated $>0.3$ or $<-0.3$ with any of the two patterns excluded from the table)

|  | "fruit, fish and vegetables" | "meat, potatoes and sweets" |
| :---: | :---: | :---: |
| Food item |  |  |
| peach | 0.55 | - |
| garlic | 0.52 | - |
| fresh vegetable/cereal soup | 0.52 | - |
| leek | 0.51 | - |
| plum | 0.51 | - |
| melon/watermelon | 0.50 | - |
| chard | 0.49 | - |
| carrot | 0.49 | - |
| onion | 0.49 | - |
| fresh white fish | 0.49 | - |
| herbs | 0.48 | - |
| cauliflower | 0.48 | - |
| pear | 0.48 | - |
| legumes, any | 0.46 | - |
| pumpkin | 0.46 | - |
| cabbage | 0.46 | - |
| nectarine | 0.46 | - |
| orange | 0.46 | - |
| courgette | 0.45 | - |
| broccoli | 0.45 | - |
| fresh fruit | 0.45 | - |
| grape | 0.45 | - |
| mango | 0.44 | - |
| apricot | 0.44 | - |
| lettuce | 0.43 | - |
| tomato | 0.43 | - |
| turnip | 0.43 | - |
| chickpeas | 0.42 | - |
| apple | 0.42 | - |
| cherry | 0.42 | - |
| pineapple | 0.42 | - |
| spinach | 0.41 | - |
| small game | 0.41 | - |
| olive oil | 0.40 | - |
| white | 0.39 | - |
| kiwi | 0.39 | - |
| fresh fatty fish | 0.39 | - |
| mandarin/tangerine | 0.38 | - |
| fava beans | 0.37 | - |
| celery | 0.37 | 0.31 |


| brussels spouts | 0.37 | - |
| :--- | :---: | :---: |
| forest fruits | 0.37 | - |
| lemon | 0.37 | - |
| aubergine | 0.36 | - |
| turkey | 0.36 | - |
| kidney | 0.35 | - |
| french beans | 0.35 | - |
| chicken boiled | 0.35 | - |
| vegetables, any | 0.34 | - |
| olive | 0.34 | - |
| smoked white fish | 0.34 | - |
| veal | 0.33 | - |
| sweet peppers | 0.32 | - |
| bitter melon | 0.32 | - |
| stuffed vegetables | 0.32 | 0.44 |
| potato tortilla | 0.32 | 0.31 |
| any poultry | 0.31 | - |
| chicken | 0.31 | 0.36 |
| parsnip | 0.30 | - |
| beetroot | 0.30 | - |
| Danish pastries | - | 0.46 |
| sweet rolls | - | 0.45 |
| custard cream | - | 0.44 |
| potatoes(boiled/mashed) | - | 0.44 |
| cakes | - | 0.42 |
| potato dumpling | - | 0.39 |
| meat stew | - | 0.39 |
| beef burger | - | 0.38 |
| total sweets or bonbons | - | 0.37 |
| canned fruit | - | 0.37 |
| halva | - | 0.31 |
| boiled sweets/toffees | - | 0.36 |
| pickled vegetables | - | 0.35 |
| smoked game | - | 0.35 |
| puddings | - | 0.35 |
| chocolate, any | - | 0.34 |
| cottage cheese | - | 0.34 |
| radish | - | 0.34 |
| bacon | - | 0.33 |
| eggs, any | - | 0.33 |
| table sugar | - | 0.33 |
| potato salad | - |  |
| doughnuts/other pastries | chocolate (plain) | - |
| butter, any |  |  |
| frankfurter | - | -32 |
|  | - | - |

egg-based desserts
0.31
ketchup
0.30

Figure 6.3.2.1.1. Scree plot for overall data combined


Figure 6.3.2.1.2. Scree plots for each country separately




Figure 6.3.2.1.3. Mean individual "meat, potatoes and sweets" and "fruit, fish and vegetables" pattern intake for each country separately.


### 6.3.2.2 For multicentre data

The scree plot from the PCA (figure 6.3.2.2.1) showed a break in the curve after two or six or nine components. We derived two principal components since they provided a more meaningful interpretation. Table 6.3.2.2.1 shows how individual foods were correlated with each of these patterns at ten countries. This table shows little similarity between individual's patterns from country to country - that is, the patterns mean different things in different countries. This makes interpretation of the results more difficult, and is another way to represent the heterogeneity between countries because of the different local diets

Additionally, according to the table, at more than eight countries, the first pattern was characterized by high consumption of brown wholemeal bread, vegetables and fruits and fresh fatty fish and the second one was closely associated with intakes of white bread, cakes , muffins, any butter, chips/fries, beef and sausages, eggs, mayonnaise and crème fraiche. For simplicity purposes we labelled the first pattern as "fruit, fish and vegetables" and the second one as "meat potatoes and sweets". However, a closer investigation of the "fruit,
fish and vegetables" pattern reveals food items that are not only "fish, fruit and vegetables" or "meat, potatoes and sweets". For example, in Table 6.3.2.2.1 a "fish, fruit and vegetables" pattern correlated highly with chicken in stews (Belgium, Finland), lamb (Belgium, South Germany, Sweden), tinned fatty fish (Belgium, Denmark, UK), turkey roast (Belgium, Finland, South Germany, Portugal, Poland), veal (Belgium, Finland), smoked game (Belgium, Finland, Portugal, Poland), sausages (Belgium), fresh fatty fish (Belgium, Denmark, Finland, North Germany, Portugal, UK and Sweden), full fat butter (Denmark), mayonnaise (Finland), pizza (North Germany), butter, any (Finland), crème fraiche (Finland, North Germany, South Germany, Poland, UK), meat pies (Finland, Poland), cured pork (Finland), frankfurter (Finland, Poland), bacon cubes (Finland) and smoked lamb (Finland, South Germany).

Figure 6.3.2.2.1. Scree plot for multicentre data


Table 6.3.2.2.1. How correlation coefficients of food items with identified dietary patterns vary between countries*

|  | Belgium (Ghent) |  | Denmark (Odense ) |  | Finland (Helsinki) |  | North Germany Berlin( |  | South Germany Munich |  | Holland (Amsterda m) |  | Portugal (Coimbra) |  | Poland <br> (Lodz) <br> (Katowice) |  | UK <br> (London) (Southampto <br> n) |  | Sweden(Karolinska)(Gothenburg)(Umea)(Uppsala) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | I** | II*** | I | II | I | II | I | II | I | II | I | II | I | II | I | II | I | II | I | II |
| broccoli | 0.63 | - | 0.53 | - | 0.48 | - | 0.46 | - | - | - | 0.39 | - | 0.35 | - | 0.38 | - | 0.34 | - | 0.31 | - |
| tomato | 0.61 | - | 0.61 | - | - | - | 0.54 | - | 0.53 | - | 0.55 | - | 0.45 | - | 0.47 | - | 0.44 | - | 0.34 | - |
| pumpkin | 0.60 | - | - | - | 0.46 | - | 0.56 | - | 0.32 | - | 0.35 | - | 0.42 | - | 0.38 | - | 0.39 | - | - | - |
| radish | 0.59 | - | 0.40 | - | - | - | - | - | - | - | 0.35 | - | 0.44 | - | 0.42 | - | - | - | 0.33 | - |
| pineapple | 0.59 | - | 0.44 | - | 0.36 | - | 0.51 | - | 0.61 | 0.35 | 0.47 | - | 0.51 | - | 0.41 | - | - | - | 0.31 | - |
| cucumber | 0.58 | - | 0.44 | - | - | - | 0.35 | - | 0.58 | - | 0.39 | - | 0.32 | - | 0.31 | - | 0.35 | - | - | - |
| cauliflower | 0.58 | - | 0.43 | - | 0.63 | 0.39 | 0.38 | 0.30 | - | - | 0.33 | - | 0.35 | - | 0.49 | - | - | - | 0.50 | - |
| coleslaw | 0.57 | - | 0.31 | - | 0.59 | 0.48 | 0.34 | 0.35 | - | - | 0.48 | - | - | - | 0.32 | - | - | - | - | - |
| onion | 0.57 | - | 0.53 | - | 0.38 | - | 0.39 | - | 0.63 | 0.32 | 0.41 | - | 0.47 | - | 0.47 | - | 0.50 | - | 0.34 | - |
| caper | 0.54 | - | - | - | 0.50 | - | 0.37 | 0.39 | - | - | 0.32 | - | - | - | - | - | - | - | - | - |
| white bread | 0.54 | 0.37 | - | 0.37 | 0.52 | 0.74 | - | 0.33 | - | - | - | - | - | 0.54 | 0.37 | 0.32 | - | - | - | 0.36 |
| chickpeas | 0.52 | - | - | - | 0.34 | - | - | - | - | - | - | - | - | - | - | - | 0.42 | - | 0.50 | - |
| chicken in stews | 0.52 | - | - | - | 0.30 | 0.36 | - | 0.33 | - | - | - | - | - | - | - | - | - | - | - | - |
| pear | 0.51 | - | 0.43 | - | 0.49 | 0.40 | - | - | 0.49 | - | - | - | 0.32 | - | 0.41 | - | - | - | 0.34 | - |
| banana | 0.51 | 0.37 | 0.47 | - | 0.30 | - | - | - | 0.36 | - | - | - | - | - | - | - | - | - | - | - |
| melon watermelon | 0.51 | 0.32 | - | - | 0.42 | - | - | - | 0.61 | 0.34 | 0.59 | - | 0.46 | - | - | - | - | - | 0.34 | - |
| lamb | 0.51 | - | - | 0.36 | - | - | - | 0.37 | 0.35 | - | - | - | - | - | - | - | - | - | 0.33 | - |
| turkey roast | 0.51 | - | - | - | 0.48 | 0.45 | - | - | 0.53 | - | - | - | 0.50 | - | 0.34 | - | - | - | - | - |
| tinned fatty fish | 0.51 | - | 0.34 | - | - | 0.34 | - | - | - | - | - | - | - | 0.33 | - | - | 0.32 | - | - | - |
| Brussels sprouts | 0.50 | - | 0.42 | - | 0.38 | - | 0.42 | - | 0.38 | - | 0.37 | - | - | - | 0.50 | 0.30 | - | - | 0.36 | - |
| rhubarb | 0.50 | - | - | - | 0.48 | 0.38 | 0.37 | - | - | - | - | - | - | - | - | - | - | - | - | - |
| veal | 0.50 | - | - | 0.34 | 0.41 | 0.66 | - | 0.44 | - | - | - | 0.39 | - | 0.52 | - | 0.48 | - | - | - | - |
| carrot | 0.49 | - | 0.41 | - | 0.46 | 0.32 | - | - | 0.63 | - | 0.47 | - | 0.43 | - | 0.45 | - | - | - | 0.37 | - |
| smoked game | 0.49 | 0.32 | - | - | 0.55 | 0.69 | - | - | - | - | - | - | 0.30 | - | 0.39 | 0.36 | - | - | - | - |

Analysis of G.A ${ }^{2}$.L.E.N data: effect of diet on asthma

| celery | 0.48 | - | 0.49 | - | 0.54 | - | 0.35 | - | 0.44 | - | 0.65 | - | - | - | 0.48 | - | 0.31 | - | 0.51 | - |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| nectarine | 0.48 | - | 0.35 | - | 0.50 | 0.33 | 0.50 | - | 0.44 | 0.30 | 0.33 | - | 0.38 | - | 0.51 | - | 0.31 | - | - | - |
| sausages | 0.48 | - | - | 0.32 | - | 0.34 | - | 0.30 | - | - | - | - | - | 0.40 | - | 0.48 | - | 0.45 | - | 0.30 |
| any legumes | 0.47 | - | - | - | 0.43 | - | - | - | - | - | - | - | - | - | - | - | - | - | 0.52 | - |
| French beans | 0.47 | - | - | - | - | - | 0.52 | - | - | - | 0.44 | - | 0.33 | 0.35 | 0.40 | - | - | - | 0.36 | - |
| artichoke | 0.47 | - | - | - | - | - | 0.37 | - | - | - | - | - | - | - | - | - | - | - | 0.44 | - |
| apricot | 0.46 | 0.33 | 0.49 | - | - | - | 0.52 | - | 0.66 | 0.40 | 0.68 | - | 0.33 | - | 0.51 | - | - | - | 0.37 | - |
| parsnip | 0.44 | - | 0.45 | - | 0.42 | - | - | - | - | - | 0.33 | - | - | - | - | - | - | - | 0.43 | - |
| herbs | 0.43 | - | 0.54 | - | 0.56 | - | 0.50 | - | 0.70 | 0.38 | 0.33 | - | 0.53 | - | 0.62 | 0.32 | 0.59 | - | 0.62 | - |
| peach | 0.43 | - | 0.47 | - | 0.69 | 0.67 | 0.54 | - | - | - | - | - | 0.36 | - | 0.39 | - | 0.50 | - | 0.36 | - |
| fresh fatty fish | 0.43 | - | 0.30 | - | 0.46 | 0.49 | 0.44 | - | - | - | - | - | 0.41 | - | - | - | 0.30 | - | 0.41 | - |
| fresh white fish fresh vegetable/cereal | 0.43 | - | 0.45 | - | - | - | 0.34 | - | - | - | - | - | - | - | 0.30 | - | - | - | 0.33 | - |
| soup | 0.43 | - | - | - | 0.51 | 0.48 | - | - | 0.34 | 0.38 | - | - | 0.39 | - | 0.34 | 0.36 | 0.34 | - | 0.43 | - |
| leek | 0.42 | - | 0.62 | - | 0.66 | - | 0.62 | - | 0.35 | - | 0.61 | - | 0.62 | - | - | - | 0.38 | - | 0.45 | - |
| olive | 0.42 | - | 0.62 | - | 0.54 | 0.43 | 0.35 | - | 0.59 | - | 0.40 | - | - | 0.31 | 0.36 | - | - | - | 0.59 | - |
| ice cream | 0.42 | - | - | - | 0.43 | 0.57 | - | - | - | - | - | - | - | 0.40 | - | 0.40 | - | 0.31 | - | 0.32 |
| dressing sauces | 0.42 | - | - | - | - | - | - | - | - | - | - | 0.44 | - | - | - | - | 0.49 | - | - | 0.31 |
| marmalade | 0.41 | 0.34 | - | - | 0.31 | - | - | - | - | - | - | - | - | - | 0.42 | 0.49 | - | - | - | - |
| lettuce | 0.41 | - | 0.52 | - | 0.32 | - | 0.41 | - | 0.45 | - | 0.42 | - | - | - | - | - | 0.52 | - | 0.39 | - |
| mango | 0.41 | 0.35 | 0.45 | - | 0.54 | - | 0.56 | - | - | - | 0.57 | - | 0.55 | - | 0.42 | - | 0.46 | - | 0.40 | - |
| grapefruit | 0.41 | - | 0.36 | - | 0.38 | - | 0.31 | - | 0.65 | 0.32 | 0.47 | - | - | - | 0.44 | - | - | - | 0.32 | - |
| wild greens | 0.40 | 0.45 | 0.30 | - | - | - | - | - | - | - | - | - | - | - | - | - | - | - | - | - |
| turnip | 0.40 | 0.41 | 0.38 | - | - | - | - | - | - | - | 0.35 | - | 0.39 | - | 0.45 | - | - | - | - | - |
| white sauce | 0.40 | 0.33 | 0.42 | - | - | 0.38 | 0.40 | 0.38 | - | 0.34 | 0.40 | - | - | - | 0.32 | - | - | - | 0.42 | - |
| hard cheese | 0.40 | - | - | - | - | - | - | - | 0.32 | - | 0.31 | - | - | - | - | 0.50 | - | - | - | - |
| sour milk | 0.39 | - | - | - | 0.40 | 0.49 | - | - | - | - | - | - | - | - | - | - | - | - | - | - |
| Greek yoghurt | 0.39 | - | - | - | 0.38 | 0.43 | - | - | - | - | - | - | - | - | - | - | 0.31 | - | - | - |
| squeezed fresh fruit | 0.38 | - | 0.49 | - | 0.48 | - | 0.42 | - | 0.42 | - | 0.34 | - | 0.41 | - | 0.32 | - | 0.45 | - | 0.42 | - |
| canned fruit | 0.38 | - | - | - | - | - | - | - | 0.33 | - | - | - | - | - | 0.44 | 0.53 | - | - | - | - |

Analysis of G.A ${ }^{2}$.L.E.N data: effect of diet on asthma


Analysis of G.A ${ }^{2}$.L.E.N data: effect of diet on asthma


Analysis of G.A ${ }^{2}$.L.E.N data: effect of diet on asthma

| pudding and desserts | - | - | - | - | - | 0.33 | - | 0.41 | - | - | - | - | 0.34 | 0.32 | - | 0.30 | - | - | - |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| pancakes | - | - | - | 0.36 | - | 0.39 | - | 0.49 | - | - | - | 0.42 | - | - | - | - | - | - | - |
| sweet biscuits | - | - | - | - | - | - | - | - | - | - | - | - | - | - | 0.31 | 0.32 | - | - | - |
| crisps, fries | - | - | - | - | - | - | - | - | - | - | 0.40 | - | - | 0.32 | - | 0.35 | - | - | - |
| thin biscuits | - | - | - | - | - | - | 0.33 | - | - | - | - | - | - | 0.31 | 0.34 | - | - | - | - |
| sweet biscuits | - | - | - | - | - | 0.32 | - | - | - | - | - | - | - | 0.59 | 0.35 | 0.43 | - | 0.42 | - |
| rice, any | - | - | 0.39 | - | 0.32 | 0.32 | - | - | - | - | - | - | - | 0.46 | - | - | - | - | - |
| white rice | - | - | - | - | - | - | - | 0.35 | - | - | - | - | - | 0.43 | - | - | - | - | - |
| brown rice | - | - | 0.49 | - | 0.54 | 0.40 | - | - | - | - | - | - | 0.35 | - | 0.44 | 0.40 | - | - | - |
| rice noodles | - | - | - | - | - | - | - | 0.31 | - | - | 0.31 | - | - | - | 0.53 | 0.54 | - | 0.31 | - |
| jam | - | - | - | - | - | 0.32 | - | - | - | - | - | - | - | - | - | 0.33 | - | - | - |
| honey | - | - | - | - | - | - | 0.34 | - | - | - | - | - | - | - | 0.33 | - | - | - | - |
| syrup spreads | - | - | - | - | - | - | - | - | - | - | - | 0.33 | - | 0.40 | 0.45 | 0.48 | - | - | - |
| apple spread | - | - | - | - | 0.39 | - | - | - | - | - | - | - | - | - | - | 0.38 | - | 0.36 | - |
| cereal bars | - | - | - | - | - | - | - | - | - | - | - | - | - | - | - | - | - | 0.33 | - |
| halva | - | - | - | - | - | - | - | - | - | - | - | - | - | - | - | - | - | 0.32 | - |
| water ice | - | - | - | - | - | - | - | 0.45 | - | - | - | - | - | - | - | - | - | - | - |
| chocolate snack bars | - | - | - | 0.37 | - | - | - | 0.44 | - | - | - | 0.38 | - | - | - | 0.38 | - | 0.42 | - |
| chocolate (plain) | - | - | - | - | - | - | - | - | - | - | - | - | - | 0.39 | - | - | - | - | - |
| vegetable oil | - | - | 0.43 | - | - | - | 0.31 | - | 0.33 | - | - | - | - | 0.35 | - | - | - | 0.49 | 0.30 |
| sunflower oil | - | - | - | - | - | - | 0.31 | - | 0.36 | 0.41 | - | - | - | - | - | - | - | 0.47 | - |
| olive oil | - | - | 0.45 | - | 0.34 | - | 0.44 | - | 0.43 | - | - | - | - | - | - | - | 0.42 | - | 0.43 |
| full fat butter | - | - | - | 0.33 | 0.33 | - | - | 0.39 | - | - | - | - | - | 0.36 | - | - | - | 0.38 | - |
| any nuts | - | - | 0.42 | - | - | - | 0.39 | - | - | - | 0.38 | - | 0.36 | - | - | - | - | - | 0.46 |
| peanuts | - | - | - | - | - | - | - | - | - | - | - | - | - | - | - | - | - | - | 0.45 |
| cashew nuts | - | - | 0.44 | - | - | - | 0.31 | - | - | - | - | - | - | - | - | - | 0.30 | - | 0.42 |
| nut based spread nuts kidney(red), black beans | - | - - | - | - | - 0.47 | - | - | - | - | 0.30 - | - | - | - | - | 0.30 | - | - | - | ${ }^{-}$ |
| lentils | - | - | - | - | 0.37 | - | - | - | - | - | 0.52 | - | - | - | - | - | 0.48 | - | 0.51 |

Analysis of G.A ${ }^{2}$.L.E.N data: effect of diet on asthma


Analysis of G.A ${ }^{2}$.L.E.N data: effect of diet on asthma
prune
concentrated juice
(sugar)
carbonated drinks
tap water
mineral water
soda (sugar)
black tea
coffee
herbal tea
red wine
rose wine
red meat, any
beef burger
minced meat in sauce
meat stew pork cutlet, chop, steak, fillet
meat pies
meat spreads e.g. rillets
small game cured pork
salami/ gammon/ham
frankfurter
bacon, bacon cubes smoked lamb
poultry, any chicken in stews/breadcrumbs/pie s turkey in stews breadcrumbs/breadcru


Analysis of G.A ${ }^{2}$.L.E.N data: effect of diet on asthma
mbs/pies
any smoked/cured poultry
liver/ pates other offal (tongue/brain etc)
fish, any
other fish
fresh crustaceans and molluscs cured / smoked fatty fish cured or smoked white fish
tinned crustaceans eggs (fried/poached e.t.c)
milk, any
full fat milk
semi -skimmed milk
skimmed milk
yogurt, any
soya milk, any viili yoghurt like fermented milk
tofu
cheese, any
soft cheese semi hard cheese and hard cheeses(gouda emental e.t.c)
cottage cheese semi hard and hard Greek cheese
fresh cheeses (e.g. mozzarella, feta e.t.c)


Analysis of G.A ${ }^{2}$.L.E.N data: effect of diet on asthma


* Values are Pearson Correlation Coefficients between a food item and an identified dietary pattern. For clarity only food items that were correlated $>0.30$ or $<-0.30$ with a dietary pattern for each country were included in the table ${ }^{* *}$ I: fruit, fish and vegetables pattern, ${ }^{* * *}$ II: meat, potatoes and sweets pattern


### 6.3.3 Dietary patterns and respiratory outcomes

Forest plots from a random effect meta-analysis (Fig. 6.3.3.1) shows no overall statistical significant evidence that a "fruit, fish and vegetables" pattern was associated with asthma ( $\mathrm{P}=0.13$ ), atopy ( $\mathrm{P}=0.62$ ) chronic sinusitis ( $\mathrm{P}=0.89$ ), allergic rhinitis ( $\mathrm{P}=0.77$ ) or eczema $(\mathrm{P}=0.17)$, and similarly no evidence that the "meat, potatoes and sweets" pattern was associated with asthma ( $\mathrm{P}=0.20$ ), atopy $(\mathrm{P}=0.31)$, chronic sinusitis $(\mathrm{P}=0.19)$, allergic rhinitis ( $\mathrm{P}=0.30$ ) or eczema ( $\mathrm{P}=0.98$ ).

There was evidence of heterogeneity on the association of allergic rhinitis and eczema with the "fruit, fish and vegetables" pattern $\left(\mathrm{I}^{2}=69.9 \% ; \mathrm{P}<0.001\right.$ and $\mathrm{I}^{2}=58.6 \% \% ; \mathrm{P}=0.01$ respectively). In this case, it was evident from a visual inspection that the two German centres had qualitatively different results to the others, differences which were statistically significant; in fact, when countries were analysed separately, increased "fruit, fish and vegetables" pattern intake was associated with a decreased risk of allergic rhinitis in North Germany and South Germany (OR per quintile $=0.56 ; 95 \%$ CI $0.35,0.84$, OR per quintile $=0.74 ; 95 \%$ CI $0.57,0.96)$ and an increased risk in Sweden and Belgium (OR per quintile $=$ $1.22 ; 95 \%$ CI $1.05,1.42$, OR per quintile $=1.52 ; 95 \%$ CI 1.02, 2.27). A visual inspection of the forest plot of the association of eczema and "fruit, fish and vegetables" pattern showed that the Amsterdam centre had significantly different results than the other centres (OR per quintile $=1.77 ; 95 \%$ CI 1.29, 2.43).

There was evidence that a "meat, potatoes and sweets" pattern was statistically significant associated with chronic sinusitis in UK, South Germany and Sweden. Furthermore a "fruit, fish and vegetables" pattern was negatively associated in South Germany and North Germany and positively associated in Belgium with allergic rhinitis. A "fruit, fish and vegetables" pattern increased the risk and a "meat, potatoes and sweets" pattern decreased the risk of eczema in Holland (Figure 6.3.3.1).

Figure 6.3.3.1. Associations between the two dietary patterns and respiratory outcomes: results of meta-analyses. OR: odds ratio.
asthma

| country | obs |  | OR ( $95 \% \mathrm{Cl}$ ) |
| :---: | :---: | :---: | :---: |
| Belgium | 146 | $\cdots$ | 0.88 (0.57, 1.36) |
| Denmark | 354 | $\square$ | 1.09 (0.81, 1.47) |
| Finland | 155 | T | 1.02 (0.73, 1.43) |
| North Germany | 177 | . | 0.93 (0.59, 1.46) |
| South Germany | 194 | - | 1.13 (0.82, 1.56) |
| Holland | 215 | 1. | 1.29 (0.88, 1.90) |
| Portugal | 259 | $\square$ | 1.37 (0.99, 1.90) |
| Poland | 210 | $\rightarrow$ | 0.90 (0.65, 1.26) |
| UK | 171 | $\cdots$ | 1.36 (0.94, 1.97) |
| Sweden | 1176 | $\rightarrow$ | 1.02 (0.91, 1.16) |
| Overall (l-square | d $0.0 \%, p=0.579$ ) | 0 | 1.07 (0.98, 1.16) |

chronic sinusitis

| country | obs |  | OR (95\% CI) |
| :---: | :---: | :---: | :---: |
| Belgium | 146 | $\ldots$ | 1.01 (0.67, 1.53) |
| Denmark | 354 | $\square$ | 1.16 (0.79, 1.72) |
| Finland | 155 |  | 1.07 (0.53, 2.19) |
| North Germany | 177 | [ | 0.99 (0.62, 1.57) |
| South Germany | 194 | $\square$ | 1.22 (0.84, 1.76) |
| Holland | 215 | $\square$ | 1.33 (0.94, 1.89) |
| Portugal | 259 | - | 1.15 (0.82, 1.61) |
| Poland | 210 | $\cdots$ | 0.84 (0.59, 1.18) |
| UK | 171 | T | 0.97 (0.63, 1.51) |
| Sweden | 1176 | $\cdots$ | 0.91 (0.77, 1.06) |
| Overall ( 1 -squared $=0.0 \%, \mathrm{p}=0.594$ ) |  | ¢ | 1.01 (0.91, 1.11) |

allergic rhinitis

eczema

| country | obs |  | OR (95\% CI) |
| :---: | :---: | :---: | :---: |
| Belgium | 146 | $\xrightarrow[\square]{\square}$ | 1.19 (0.85, 1.68) |
| Denmark | 353 | $\rightarrow 1$ | 0.94 (0.73, 1.22) |
| Finland | 155 | 1 | 1.02 (0.70, 1.48) |
| North Germany | 177 | $\square$ | 1.09 (0.75, 1.59) |
| South Germany | 194 | $\rightarrow$ | 0.79 (0.61, 1.01) |
| Holland | 215 | $\cdots \rightarrow$ | 1.77 (1.29, 2.43) |
| Portugal | 259 | $\rightarrow$ | 0.90 (0.67, 1.21) |
| Poland | 209 | $\stackrel{\square}{\square}$ | 1.34 (0.96, 1.87) |
| UK | 171 | $\square$ | 1.18 (0.80, 1.76) |
| Sweden | 1176 | + | 1.18 (1.02, 1.36) |
| Overall (l-square | $\mathrm{d}=58.6 \%, \mathrm{p}=0.010$ ) | Q | 1.10 (0.96, 1.27) |

asthma

chronic sinusitis


eczema

| country | obs | OR (95\% CI) |
| :---: | :---: | :---: |
| Belgium | 146 | 1.28 (0.90, 1.82) |
| Denmark | 353 | 0.83 (0.64, 1.08) |
| Finaland | 155 | 1.04 (0.74, 1.46) |
| North Germany | 177 | 0.95 (0.66, 1.39) |
| South Germany | 194 | 1.18 (0.92, 1.52) |
| Holland | 215 | 0.67 (0.50, 0.89) |
| Portugal | 259 | 1.15 (0.88, 1.51) |
| Poland | 209 | 1.08 (0.77, 1.50) |
| UK | 171 | 0.85 (0.59, 1.20) |
| Sweden | 1176 | 1.07 (0.93, 1.24) |
| Overall (l-squared $=43.6 \%, \mathrm{p}=0.068$ ) |  | 1.00 (0.89, 1.13) |

Analysis of G.A ${ }^{2}$.L.E.N data: effect of diet on asthma
atopy
fish, fruit and vegetables pattern

atopy
meat, potatoes and sweets pattern

| country | obs |  | OR (95\% Cl) |
| :---: | :---: | :---: | :---: |
| Belgium | 135 | $4$ | 1.03 (0.74, 1.44) |
| Denmark | 332 | $\rightarrow$ | 0.92 (0.70, 1.21) |
| North Germany | 176 |  | 1.04 (0.69, 1.55) |
| South Germany | 186 |  | 1.04 (0.80, 1.35) |
| Holland | 201 | , | 0.96 (0.71, 1.31) |
| Portugal | 255 | - | 1.21 (0.84, 1.75) |
| Poland | 159 | + | 1.11 (0.80, 1.54) |
| UK | 153 | 1 | 1.02 (0.72, 1.45) |
| Sweden | 1140 | 5 | 1.08 (0.93, 1.24) |
| Overall ( 1 -squared $=0.0 \%, \mathrm{p}=0.980$ ) |  | $\phi$ | 1.05 (0.96, 1.14) |

### 6.3.4 Exhaustive single food analysis and respiratory outcomes

The results from our two step exhaustive single food analysis where each of the 238 individual food items (xilopites was omitted due to limited observations) was separately entered into a random intercept logistic analysis are presented in smile plots in Fig. 6.3.4.1. To be considered significant with the false discovery rate controlled at $20 \%$, an individual result here needs to be above the horizontal dashed line. Odds ratios of food items that were above the $20 \%$ threshold are shown in tables 6.3.4.1-6.3.4.5 for the analysis not adjusted for other foods at the first step (but adjusted for age, sex, body mass index and smoking status) and the adjusted analysis at the second step (adjusted for age, sex, body mass index and smoking status and for foods declared significant at the first step). Furthermore, our analysis was stratified by atopic status. An additional two-step ESFA was performed for the overall sample of adults controlling the FDR at $5 \%$.

Overall, there was evidence of an association between individual food items and asthma, after taking proper account of the number of comparison- that is we set our false discovery rate to be at $20 \%$. Specifically, for the adjusted analysis at the second step a number of vegetables (legumes, chard, lemon, radish), cherries, cured or smoked fatty fish, condensed milk, ice-cream, beans (kidney, red and black, wholemeal bread and pasta seem to decrease the risk of asthma, while couscous, turnip, ketchup, normal margarine, nut based spreads, other fish seafood, and thin biscuits (knackerbrod) seems to increase the risk (Table 6.3.4.1). Condensed milk, cured or smoked fatty fish, turnip and couscous remained statistically significantly associated with asthma, even when we stratified our analysis by non atopics and atopics or when we controlled the FDR for the overall sample at 5\% (Table 6.3.4.1).

Furthermore, in our two-step ESFA (allowing the FDR at 20\%) plant, vegetables and fruits (apple, chickpeas, fenugreek and cabbage), fresh white fish, raisins and sultan, sour cream, and meat stew were negatively associated with chronic sinusitis. Positive associations with chronic sinusitis were observed with table sugar, okra, pumpkin, turkey roasted or boiled, soya beans, and potatoes (mashed/boiled) (Table 6.3.4.2). Cabbage and Vili yogurt like fermented milk, okra and pumpkin remained significant in the majority of our ESFA models when we stratified for atopics, non-atopics and when we controlled the FDR of our ESFA at $5 \%$.

Several food items such as fruits (peach, kiwi) butter, peanuts, beef burger, cured or smoked fatty fish, Greek Style yogurt, smoked game and tofu seems to be protective for allergic rhinitis (6.3.4.3). Condensed milk, rhubarb and lentils, bitter melon, crisp fried cakes, Greek style yoghurt were inversely associated with eczema (Table 6.3.4.4) for our two-step ESFA model (allowing the FDR at 20\%). In the majority of the models (subgroup analysis for atopy and allowing the rate of our false discoveries to be at $5 \%$ ), protective associations between peach, peanuts, and tofu and Allergic Rhinitis remained statistical significant. Similar results were observed between Bitter Mellon, Greek Style yogurt, smoked game and eczema.

Finally, atopy was negatively associated with smoked game, pancakes, parsnip, crisp fried cakes, Double clotted cream, moussaka, cured pork and positively associated with soft cheeses, legumes, white rice, beer, sweet potato, hot/cold/roast beef, broccoli, thin biscuits (knackerbrod), total sweets and couscous in our two-step ESFA (6.3.4.5). Crisp fried cakes, moussaka, beer, thin biscuits, hot/cold/roast beef remained statistical significant after controlling our rate of false discoveries at $5 \%$.

Table 6.3.4.1. Associations between food items and asthma; odds ratio (OR) and P-values for statistical significant foods at first step and second step of Exhaustive Single Food Analysis (ESFA); ESFA (1 ${ }^{\text {st }}$ step) controlling the False Discovery Rate (FDR) at 20\%; ESFA (2 ${ }^{\text {nd }}$ step) controlling the FDR at $20 \%$; ESFA ( $2^{\text {nd }}$ step) controlling the FDR at $20 \%$ for non-atopics; ESFA ( $2^{\text {nd }}$ step) controlling the FDR at $20 \%$ for atopics; ESFA ( $2{ }^{\text {nd }}$ step) controlling the FDR at $5 \%$.

| $\begin{array}{r} \text { ESFA }\left(\mathbf{1}^{\text {st }}\right. \text { st } \\ \text { FDR=20 } \\ \text { all p-values< } \end{array}$ |  | $\begin{gathered} \text { ESFA }\left(2^{\text {nd }} \text { step }\right)^{* *} \\ \text { FDR }=20 \% \\ \text { all p-values }<0.010 \\ \hline \end{gathered}$ |  | $\begin{gathered} \text { ESFA ( } \left.2^{\text {nd }} \text { step) }\right)^{* *} \\ \text { non-atopics } \\ \text { FDR=20\% } \\ \text { all p-values }<0.010 \\ \hline \end{gathered}$ |  | $\begin{gathered} \text { ESFA (2 } \left.{ }^{\text {nd }} \text { step) }\right)^{* *} \\ \text { atopics } \\ \text { FDR }=20 \% \\ \text { all p-values }<0.015 \\ \hline \end{gathered}$ |  | $\begin{gathered} \text { ESFA }\left(2^{\text {nd }} \text { step }\right)^{* *} \\ \text { FDR }=5 \% \\ \text { all p-values }<0.001 \\ \hline \end{gathered}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| food | OR | food | OR | food | OR | food | OR | food | OR |
| cakes, any | 0.79 | refined pasta | 0.81 | cakes, any | 0.70 | kidney and black beans | 0.72 | kidney (black beans) | 0.85 |
| whole meal bread | 0.84 | lemon | 0.82 | small game | 0.78 | lemon | 0.75 | lemon cured or smoked | 0.85 |
| ice -cream | 0.84 | kidney ,black beans | 0.83 | apple spread freshly squeezed | 0.79 | refined pasta | 0.80 | fatty fish | 0.85 |
| refined pasta | 0.85 | whole meal bread cured or smoked | 0.83 | juice | 0.81 | radish | 0.83 | cherries | 0.86 |
| radish | 0.86 | fatty fish | 0.85 | grapefruit chicken boiled | 0.81 | cherries cured or smoked | 0.85 | legumes, any | 0.89 |
| condensed milk cured or smoked | 0.87 | radish | 0.85 | roasted | 0.83 | fatty fish | 0.86 | condensed milk thin biscuits | 0.89 |
| fatty fish | 0.88 | ice-cream | 0.86 | condensed milk cured or smoked | 0.85 | apricot | 0.86 | (knackerbrod) | 1.09 |
| chocolate( plain) | 0.89 | cherries | 0.87 | fatty fish | 0.88 | coffee | 0.86 | turnip | 1.13 |
| cherries | 0.90 | legumes, any | 0.88 | couscous | 1.12 | condensed milk | 0.87 | couscous | 1.17 |
| cottage cheese | 0.90 | condensed milk | 0.90 | single cream | 1.18 | chocolate plain | 0.88 |  |  |
| nuts, any | 0.91 | chard | 0.92 | turnip | 1.27 | sour milk | 0.88 |  |  |
| butter, any | 0.92 | other fish sea food thin biscuits | 1.04 | sunflower oil | 1.27 | normal margarine | 1.11 |  |  |
| fenugreek | 1.05 | (knackerbrod) | 1.12 |  |  | minced beef meat | 1.12 |  |  |
| fresh crustaceans | 1.06 | ketchup | 1.12 |  |  | fenugreek | 1.12 |  |  |


| fresh white fish | 1.06 | nut based spreads | 1.13 | double clotted cream | 1.13 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| cauliflower | 1.06 | normal margarine | 1.13 | couscous | 1.14 |
| bitter mellon cured or smoked | 1.06 | turnip | 1.14 | spirits | 1.15 |
| white fish | 1.06 | couscous | 1.21 | lentils | 1.17 |
| chard | 1.07 |  |  | mayonnaise | 1.18 |
| soya milk | 1.08 |  |  | crème fraiche | 1.19 |
| white rice thin biscuits | 1.08 |  |  | turnip | 1.20 |
| (knackerbrod) | 1.09 |  |  | half fat butter | 1.23 |
| herbs | 1.09 |  |  |  |  |
| okra | 1.10 |  |  |  |  |
| squeezed fresh fruit | 1.10 |  |  |  |  |
| olive oil | 1.10 |  |  |  |  |
| avocado <br> fresh <br> vegetable/cereal | 1.11 |  |  |  |  |
| soup | 1.12 |  |  |  |  |
| mango | 1.13 |  |  |  |  |
| onion | 1.13 |  |  |  |  |
| garlic | 1.14 |  |  |  |  |
| courgette | 1.14 |  |  |  |  |
| leek | 1.16 |  |  |  |  |
| couscous | 1.19 |  |  |  |  |
| turnip | 1.20 |  |  |  |  |

Table 6.3.4.2. Associations between food items and chronic sinusitis; odds ratio (OR) and P-Values for statistical significant foods at first step and second step of Exhaustive Single Food Analysis (ESFA) ;ESFA (1 ${ }^{\text {st }}$ step) controlling the False Discovery Rate (FDR) at 20\%; ESFA (2 ${ }^{\text {nd }}$ step) controlling the FDR at $20 \%$; ESFA ( $2^{\text {nd }}$ step) controlling the FDR at $20 \%$ for non-atopics; ESFA ( $2^{\text {nd }}$ step) controlling the FDR at $20 \%$ for atopics; ESFA ( $2^{\text {nd }} s t e p$ ) controlling the FDR at $5 \%$.

| ESFA ( ${ }^{\text {st }}$ ste FDR=20\% all p-values $<0$ |  | ESFA ( $2^{\text {nd }}$ step) $)^{* *}$ <br> FDR=20\% <br> all p-values $<0.011$ |  | $\begin{gathered} \text { ESFA ( } \left.2^{\text {nd }} \text { step) }\right)^{* *} \\ \text { non-atopics } \\ \text { FDR }=20 \% \\ \text { all p-values }<0.028 \\ \hline \end{gathered}$ |  | $\begin{gathered} \text { ESFA ( } \left.2^{\text {nd }} \text { step) }\right)^{* *} \\ \text { atopics } \\ \text { FDR }=20 \% \\ \text { all p-values }<0.019 \\ \hline \end{gathered}$ |  | ESFA ( $\mathbf{2 d x}^{\text {nd }}$ step) $)^{* *}$ FDR=5\% all p-values $<0.001$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| food | OR | food | OR | food | OR | food | OR | food | OR |
| sour cream | 0.86 | cabbage | 0.79 | rhubarb | 0.57 | raisins | 0.66 | cabbage | 0.87 |
|  |  |  |  | custard |  | pickled |  | viili yogurt like |  |
| mixed candies | 0.88 | raisin, sultan | 0.80 | cream | 0.67 | vegetables | 0.72 | fermented milk | 1.11 |
| apple | 0.88 | sour cream | 0.81 | cabbage | 0.71 | cucumber | 0.73 | okra | 1.13 |
| smoked lamb | 1.05 | apple | 0.81 | small game | 0.77 | peach | 0.74 | pumpkin | 1.14 |
| turnip | 1.06 | fresh white fish | 0.84 | celery | 0.77 | pear | 0.74 |  |  |
| chapatti bread | 1.06 | chickpeas | 0.85 | bacon <br> cubes | 0.78 | fresh fatty fish | 0.75 |  |  |
|  |  |  |  | legumes |  |  |  |  |  |
| cauliflower | 1.10 | fenugreek | 0.88 | ,any | 0.80 | cabbage | 0.76 |  |  |
|  |  |  |  | fresh |  |  |  |  |  |
| turkey roasted, boiled | 1.10 | meat stew | 0.90 | cheeses | 0.80 | fava beans | 0.78 |  |  |
| viili yogurt like |  | turkey roasted, |  |  |  |  |  |  |  |
| fermented milk | 1.13 | boiled | 1.05 | plum | 0.81 | fresh white fish | 0.78 |  |  |
|  |  |  |  | cottage |  | soda without |  |  |  |
| table sugar | 1.13 | soya beans | 1.10 | cheese | 0.81 | sugar | 0.79 |  |  |
| olive | 1.14 | okra | 1.11 | apple | 0.81 | apple | 0.79 |  |  |
| aubergine | 1.14 | pumpkin | 1.11 | garlic | 0.82 | white wine | 0.79 |  |  |
| okra | 1.16 | table sugar | 1.11 | sour cream | 0.82 | sweet corn | 0.79 |  |  |
| pumpkin | 1.16 | potatoes(boiled/ mashed) | 1.16 | liver pates | 0.87 | chard | 0.80 |  |  |

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| carrot | 1.16 | salads <br> cluster | 0.87 | fresh fruit | 0.80 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| courgette | 1.16 | beans doughnuts | 1.10 | garlic | 0.80 |
| olive oil | 1.17 | buns | 1.14 | small game | 0.81 |
| white sauce | 1.24 | pumpkin | 1.15 | fig | 0.83 |
|  |  |  |  | kidney black |  |
|  |  | okra <br> single | 1.20 | beans | 0.84 |
|  |  | cream any red | 1.21 | other offal | 0.85 |
|  |  | meat chocolates | 1.22 | rice | 0.85 |
|  |  | bars | 1.24 | soya beans | 0.90 |
|  |  |  |  | viili yogurt like |  |
|  |  | white sauce | 1.27 | fermented milk | 1.07 |
|  |  |  |  | margarine, any | 1.10 |
|  |  |  |  | okra | 1.15 |
|  |  |  |  | low fat butter | 1.18 |
|  |  |  |  | cakes | 1.20 |
|  |  |  |  | mayonnaise | 1.21 |
|  |  |  |  | apple spread | 1.22 |
|  |  |  |  | juice with sugar | 1.24 |
|  |  |  |  | full fat milk | 1.25 |
|  |  |  |  | poultry, any | 1.28 |
|  |  |  |  | courgette | 1.33 |
|  |  |  |  | total sweets | 1.35 |

*Adjusted for age, sex, body mass index and smoking status. **Adjusted for age, sex, body mass index, smoking status and all the foods that were significant at the univariate analysis ( $1^{\text {st }}$ step) ${ }^{* * *}$ Bold type indicates foods which were statistically associated with chronic sinusitis across 3 or more of the different procedures being used.

Table 6.3.4.3. Associations between food items and allergic rhinitis; odds ratio (OR) and P-Values for statistical significant foods at first step and second step of Exhaustive Single Food Analysis (ESFA); ESFA (1 ${ }^{\text {st }}$ step) controlling the False Discovery Rate (FDR) at 20\%, ESFA (2 ${ }^{\text {nd }}$ step) controlling the FDR at $20 \%$, ESFA ( $2^{\text {nd }}$ step) controlling the FDR at $20 \%$ for non-atopics, ESFA ( $2^{\text {nd }}$ step) controlling the FDR at $20 \%$ for atopics, ESFA ( $2^{\text {nd }}$ step) controlling the FDR at $5 \%$.


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*Adjusted for age, sex, body mass index and smoking status. **Adjusted for age, sex, body mass index, smoking status and all the foods that were significant at the univariate analysis ( $1^{\text {st }}$ step). ${ }^{* * *}$ Bold type indicates foods which were statistically significant associated with allergic rhinitis across 3 or more of the different procedures being used.

Table 6.3.4.4. Associations between food items and Eczema; odds ratio (OR) and P-Values for statistical significant foods at first step and second step of Exhaustive Single Food Analysis (ESFA); ESFA (1 ${ }^{\text {st }}$ step) controlling the False Discovery Rate (FDR) at 20\%, ESFA (2 ${ }^{\text {nd }}$ step) controlling the FDR at $20 \%$, ESFA ( $2^{\text {nd }}$ step) controlling the FDR at $20 \%$ for non-atopics, ESFA ( $2^{\text {nd }}$ step) controlling the FDR at $20 \%$ for atopics, ESFA ( $2{ }^{\text {nd }}$ step) controlling the FDR at $5 \%$.

| $\begin{gathered} \text { ESFA }\left(1^{\text {st }} \text { step }\right)^{*} \\ \text { FDR }=20 \% \\ \text { all p-values }<0.009 \end{gathered}$ |  | ESFA ( $2^{\text {nd }}$ step) $)^{* *}$ <br> FDR=20\% <br> all $p$-values $<0.015$ |  | $\begin{gathered} \left.\hline \text { ESFA (2 } 2^{\text {nd }} \text { step) }\right)^{* *} \\ \text { Non-atopics } \\ \text { FDR }=20 \% \\ \text { all p-values }<0.014 \\ \hline \end{gathered}$ |  | $\begin{gathered} \left.\hline \text { ESFA ( } 2^{\text {nd }} \text { step) }\right)^{* *} \\ \text { atopics } \\ \text { FDR }=20 \% \\ \text { all p-values }<0.015 \end{gathered}$ |  | ESFA ( $2^{\text {nd }}$ step) $)^{* *}$ FDR=5\% all $p$-values $<\mathbf{0 . 0 0 1}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| food | OR | food | OR | food | OR | food | OR | food | OR |
|  |  |  |  | chips/ French |  |  |  |  |  |
| crisp fried cakes | 0.86 | rhubarb | 0.84 | fries | 0.74 | crisp fried cakes | 0.73 | rhubarb | 0.83 |
| bitter melon | 0.90 | crisp fried cakes | 0.86 | wholemeal bread | 0.77 | cottage cheese | 0.80 | crisp fried cakes | 0.84 |
| pizza | 0.90 | bitter melon | 0.89 | rhubarb | 0.79 | fresh white fish | 0.83 | bitter melon | 0.88 |
| Greek style yogurt | 0.91 | Greek style yogurt | 0.90 | cashew nuts | 0.81 | pickled vegetables | 0.87 | Greek style yogurt | 0.89 |
| kneipp bread | 0.92 | lentils | 0.93 | full fat milk | 0.84 | dressing sauces | 0.87 | game, other | 0.92 |
| game, other | 0.94 | game, other | 0.94 | courgette | 0.88 | red meat, any | 0.88 | lentils | 0.93 |
| tomato | 1.10 | condensed milk | 0.94 | wholemeal pasta | 0.89 | condensed milk | 0.89 | sour cream | 1.10 |
|  |  | turkey |  |  |  |  |  |  |  |
| wholemeal bread turkey | 1.12 | roasted/boiled | 1.11 | meat pies | 0.89 | fermented milk | 0.90 |  |  |
| roasted/boiled | 1.13 | eggs all | 1.12 | chard | 1.09 | game, other | 0.90 |  |  |
| sour cream | 1.13 | sour cream | 1.12 | brown <br> wholemeal double clotted | 1.09 | chard | 1.12 |  |  |
| honey | 1.14 |  |  | cream | 1.14 |  |  |  |  |
| vegetable oil | 1.17 |  |  | casserole | 1.15 |  |  |  |  |
| olive oil | 1.19 |  |  | Greek coffee | 1.16 |  |  |  |  |

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| potatoes | 1.17 |
| :--- | :--- |
| olive oil | 1.19 |
| prune | 1.24 |
| bread, any | 1.25 |

*Adjusted for age, sex, body mass index and smoking status. **Adjusted for age, sex, body mass index, smoking status and all the foods that were significant at the univariate analysis ( $1^{\text {st }}$ step). ${ }^{* * *}$ Bold type indicates foods which were statistically significant associated with eczema across 3 or more of the different procedures being used

Table 6.3.4.5. Associations between food items and Atopy; odds ratio (OR) and P-Values for statistical significant foods at first step and second step of Exhaustive Single Food Analysis (ESFA); ESFA (1 ${ }^{\text {st }}$ step) controlling the False Discovery Rate (FDR) at $20 \%$, ESFA ( ${ }^{\text {nd }}$ step) controlling the FDR at $20 \%$; ESFA ( $2^{\text {nd }}$ step) controlling the FDR at $5 \%$.

| $\begin{gathered} \text { ESFA (1 } \left.1^{\text {st }} \text { step) }\right)^{*} \\ \text { FDR }=20 \% \\ \text { all p-values<0.015 } \end{gathered}$ |  | $\begin{gathered} \text { ESFA }\left(2^{\text {nd }} \text { step }\right)^{* *} \\ \text { FDR }=20 \% \\ \text { all p-values }<0.011 \\ \hline \end{gathered}$ | $\begin{gathered} \text { ESFA }\left(2^{\text {nd }} \text { step) }\right)^{* *} \\ \text { FDR=5\% } \\ \text { all p-values }<0.001 \end{gathered}$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| ood | OR | food | OR | food | OR |
| parsnip | 0.83 | smoked game | 0.81 | crisp fried cakes | 0.86 |
| double or clotted cream | 0.84 | parsnip | 0.85 | moussaka | 0.88 |
| pancakes | 0.85 | pancakes | 0.85 | beer | 1.10 |
| carrot | 0.86 | crisp fried cakes | 0.86 | thin biscuits (knackerbrod) hot cold roast beef/ | 1.13 |
| crisp fried cakes | 0.88 | double or clotted cream | 0.87 | boiled beef | 1.14 |
| moussaka | 0.88 | moussaka | 0.87 |  |  |
| water ice | 0.90 | cured pork | 0.91 |  |  |
| crème fraiche | 0.90 | legumes, any | 1.07 |  |  |
| urnip | 0.92 | white rice | 1.08 |  |  |
| weet potato | 1.08 | soft cheeses | 1.09 |  |  |
| oft cheeses | 1.08 | beer | 1.09 |  |  |
| hin biscuits |  |  |  |  |  |
| knackerbrod) | 1.10 | sweet potato | 1.10 |  |  |
|  |  | hot cold roast beef/ |  |  |  |
| beer | 1.11 | boiled beef | 1.11 |  |  |
| otal sweets | 1.11 | broccoli | 1.11 |  |  |
| hot cold roast beef/ boiled |  | thin biscuits |  |  |  |
| beef | 1.14 | (knackerbrod) | 1.12 |  |  |
| any red meat | 1.15 | total sweets | 1.13 |  |  |
| white rice | 1.22 | couscous | 1.20 |  |  |

[^7]Figure 6.3.4.1. Association of foods with asthma, chronic sinusitis, allergic rhinitis and eczema: smile plot showing $P$-value against standardized logistic regression coefficient for different foods. Points on the right represent positive associations; points on the left represent negative associations. To be considered significant, controlling the false discovery rate at $\mathbf{2 0} \%$, points must lie above the dotted line.


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Atopy

Crisp fried Cakes
Moussaka

Thin biscuits

Roast beef steak

Couscous
Smoked game


## 7 Discussion and Conclusion

### 7.1 Purpose

7.2 Key Points
7.3 Interpretation of the results of the Systematic Review
7.3.1 Diet and Disease associations with the use of posterior dietary patterns identified by PCA
7.3.2 Methodological considerations of dietary pattern analysis with the use of PCA
7.3.3 Translation of posterior dietary patterns identified by PCA into an intervention

### 7.3.4 Conclusion

7.4 Interpretation of results from our simulation study
7.5 Analysis of G.A ${ }^{2}$.L.E.N data: effect of diet on asthma
7.6 Conclusions
7.7 Ideas for future research and implications of these findings
"I'm astounded by people who want to 'know' the universe when it's hard enough to find your way around Chinatown."

Woody Allen

### 7.1 Purpose

The purpose of this chapter is to:

- Interpret the results of our systematic review
- Interpret the results of our simulations
- Interpret the findings from our data analysis in the G.A ${ }^{2}$.L.E.N study
- Identify and address potential limitations of our studies
- Present ideas for potential future research


### 7.2 Key points that are discussed in this chapter

- Methodological advantages of PCR that it can capture additive and interactive effects of diet on disease and it can solve the confounding issues between dietary exposures better than analysing any single food and nutrient are not critically evaluated.
- Application of Principal Components Analysis (PCA) of diet in nutritional epidemiology requires a lot of arbitrary and subjective decisions.
- Results of Principal Components Regression (PCR) studies are contradictory, inconsistent, and difficult to be interpreted and translated into an intervention.
- An Exhaustive Single Food Analysis (ESFA) is a better method at identifying combination of foods that are causally linked to disease.
- Even when a population Principal Component of diet was causally linked to disease, PCR could not outperform ESFA in identifying the combination of foods that were causally associated with disease.
- PCR is not better at detecting an additive effect of diet on disease than an ESFA.
- PCR doesn't resolve aspects of confounding and mulitcollinearity between food intakes better than an ESFA. An ESFA method adjusted for all food intakes which were significant in the unadjusted ESFA, controls rate of false discoveries to an acceptable level.
- Dietary patterns empirically derived with the use of PCA were not associated with respiratory outcomes in the G.A ${ }^{2}$.L.E.N data-set.
- An ESFA method adjusted for all food intakes which were significant in an ESFA not adjusted for other foods (but adjusted for age, sex, bmi and smoking status) provide evidence of associations between a combination of foods and our respiratory outcomes in the G.A ${ }^{2}$.L.E.N data-set.
- In the G.A ${ }^{2}$.L.E.N analysis, diet- disease associations presented by ESFA adjusted for all food intakes which were significant in an ESFA not adjusted for other foods (but adjusted for age, sex, bmi and smoking status), could not be all causal and were difficult to be interpreted. Confounding still remains an important issue in our ESFA method and needs more careful consideration.


### 7.3 Interpretation of results from our Systematic Review

As was mentioned in Chapter 2 recent results have raised questions concerning the role of diet in the aetiology in certain chronic diseases and together with the challenges of analysing diets this has led to an explosion in the use of PCA. Ultimately, as was mentioned in the introduction, whether PCA can effectively detect associations between a combination of foods and disease across diverse populations (section 7.3.1), whether it is an objective tool of analysis (section 7.3.2) and whether this can be translated into an intervention (section 7.3.3) are three important indicators of the validity of the method and hence of their utility in epidemiological and clinical research.

### 7.3.1 Diet and disease associations with the use of posterior dietary patterns identified by PCA

Cardiovascular disease, obesity and hdl-cholesterol were consistently negatively associated with a pattern loosely defined by a high consumption of fruit, fish and vegetables. On the other hand, high consumption of a pattern consisting of red meat, chips and dairy products increased the risk.

However, there were major inconsistencies in the associations between dietary patterns and disease among the majority of the studies in our systematic review. Evidence regarding the association of disease outcomes (which appeared in more than one study) and dietary patterns were inconclusive about the beneficial or harmful effects of patterns in relation to adenoma, five different types of cancers, type II diabetes, respiratory symptoms, metabolic syndrome, overall mortality and myocardial infarction. For example, breast cancer risk was inversely associated (Hirose et al., 2007, Agurs-Collins et al., 2009), not associated (Fung et al., 2005, Kroenke et al., 2005, Robinson et al., 2004), and positively associated (Murtaugh et al., 2008) with a loosely named "prudent" pattern (Table C and D).

Inconsistencies may be partly explained by differences in sex and age since dietary patterns were highly associated with them (see paragraph 3.2.9) in diverse populations (see paragraph 3.2.2) or by the fact that PCA-derived components may be not explaining as much variation as possible in important nutrients that are potentially associated with disease risk. However,
these inconsistencies could be attributed to methodological issues such as that dietary patterns identified with the use of PCA are implicit measures of lifestyle and are not reproducible.

## Dietary patterns confounded by lifestyle factors

First, as we can observe from paragraph 3.2.9, Table C and D, patterns associated with a high intake of fruit, vegetables, and fresh fish has been associated with elements of lifestyle such as vitamin and supplement use, income, social class, strenuous exercise, education, partnership status and non-smoking status which are considered to lead to a healthier life. On the other hand, individuals who consume patterns correlated highly with meat, chips, fatty and sugary foods are more likely to lead a less healthy lifestyle (e.g. non-exercisers, smokers and drinkers). So, failure of dietary pattern analysis with the use of PCA to consistently associate combination of foods with disease across diverse populations could be due to confounding and mulitcollinearity with cultural or lifestyle factors. This is also highlighted by the fact that in several studies the dietary pattern and health association was markedly attenuated by control for confounders (Kant, 2004, Ambrosini et al., 2008, Mannisto et al. 2005, Shaheen et al. 2009).

## Reproducibility

Second, there is a vague language on the labelling of dietary patterns and the food items that constitute them which could produce spurious reproducibility when reporting.

Decisions on the label appear to be equally influenced by food items and by prevailing notions of diet and disease associations, which could be different from one place to another. For example, a "prudent", "healthy", "health conscious", "heart healthy", "healthy" "healthy/Mediterranean", "vegetables, fruit and fish" or "vegetarian" pattern and generally patterns which were correlated highly with different types of fish, fruit and vegetables were reported in all PCA studies which used food items or food groups. These patterns were mainly characterised by high intakes on food items that are considered as beneficial for human health. A "Western", "Western-like", "Western-type", "processed", "junk food", "junk", "fast food" or "unhealthy" and in general patterns which were correlated highly in high fat and processed foods, red meat and chips were observed in the majority of the populations being studied and were characterised by the items which are considered to have a harmful effect on human health (Table C and D).

However, each of these so-named healthy or unhealthy patterns derived from PCA differed from each other to greater or lesser degree on the combination of foods that were singled out (Table D). For example, people who had higher intakes of a "prudent" pattern in one study had higher intakes of cruciferous and other vegetables, fruit, whole grains, low-fat dairy products, cereals, beans, fish, and poultry (Agurs-Collins et al., 2009) and in another study had a higher intake of tea, fruits, butter, vegetable oil, breakfast cereals, fish and seafood and reduced-fat products (Kesse-Guyot et al., 2009). In similar manner, a "Western" pattern was constituted of high fat dairy, eggs, processed meat, refined grains, meat, potato, fried foods (Imamura et al., 2009) but also of pasta with meat, beef taco, beef burrito, pizza, meatballs, hamburger, fried potatoes, baked potatoes, mashed potatoes, pancake, bagels in another one (Wu et al., 2009) (Table D).

One of the main reasons for these inconsistencies is that principal component loadings differ from one study to another, so different food items contribute differently to the nutrient composition of the pattern. For example, although cereals are a consistent element of a "prudent" pattern in the majority of studies in our review its principal component loading is 0.28 (Agurs-Collins et al., 2009) or 0.39 (Kesse-Guyot et al., 2009) or 0.44 (Murtaugh et al., 2007) or 0.13 (Crozier et al., 2006). Hence, contribution of cereals to the label of a "prudent" pattern depends solely on an arbitrary cut-off point of the principal component loading of a food on a pattern. More importantly, cereals contribute differently to the nutrient composition of each of these "prudent" patterns (Table D).

In conclusion, appearance of qualitatively "similar" dietary patterns in different studies could be the reason of inconsistent and potentially misleading information on associations between diet and disease across studies. Reasons for these inconsistencies may be more apparent if studies were reporting dietary patterns quantitatively and not qualitatively.

In addition, there are several methodological considerations which are raised with the use of PCA, which call for the subjective judgment of the researcher. In the next section, a detailed discussion will be provided.

### 7.3.2 Methodological considerations of dietary pattern analysis with the use of PCA

There are several methodological issues concerning principal component analysis as applied in nutritional epidemiology. Decisions made by the researchers for preparing the data before entering the PCA and when they employ PCA involve subjectivity and have an impact on the number, type and label of patterns that are derived, reported, and analyzed. Several studies in this field have provided useful commentary on the issue of subjectivity and other methodological considerations (Newby \& Tucker, 2004, McCann et al., 2001b, Slattery \& Boucher, 1998). In this paragraph, we will discuss the results from these studies along with the results from our systematic review of Chapter 3, and provide some suggestions of how to apply PCA in nutritional epidemiology.

## Preparation of data before entering the PCA

Specifically, as presented in section 3.2.3 several studies decided to group the number of food items from the Food Frequency Questionnaire being used into predefined food groups before beginning the pattern analysis. Some of the food items were of such limited number (e.g. < 20) (Takaoka \& Norback, 2008, Igbal et al., 2008) that further collapsing was not necessary. As in Newby (2004), in our systematic review, food grouping schemes depended on a priori knowledge and on the research hypothesis under investigation.

Mc Cann (2001) (McCann et al., 2001a) and Costacou (2003) (Costacou et al., 2003) stated that a smaller number of input variables included in the PCA procedure could affect the principal component solution and explain a greater percentage of the variance intake compared with a larger number of input variables in the same study. However, this conclusion is incorrect. This was highlighted from our systematic review and Newby (2004), since studies didn't provide different results on the number of patterns and even more importantly to the total variance being explained when they grouped food items before the application of PCA (Table B).

Specifically, if you aggregate, you inevitably explain a smaller percentage of the original variance with the same number of dietary patterns. McCann found that the percentage of variance in the aggregated variables was greater, but the problem is that by aggregating you throw away a percentage of the original variance. Furthermore, in this paper, the authors found that although the number and type of dietary patterns did not change the relationship
between the dietary patterns and cancer risk was substantially attenuated when using broad food groups, suggesting that greater detail in food groupings is important. Therefore retaining a large number of foods and reducing the number of subjective decisions required seems preferable especially since the multivariate statistical approaches are actually data reduction techniques. Few studies that collapsed food items to food groups identified patterns that explained a percentage of total variance much higher ( $>55 \%$ ) above the median (median: 24\%, IQR:19.9-31.3) (Ambrosini et al., 2008, Romaguera et al., 2008, Panagiotakos et al., 2007b). So, if there is not a strong clinical hypothesis to group the food items, doesn't seem to be the need to collapse food items to food groups before entering the PCA.

In Table B, the scale of the FFQ data had a range between 5 and 11 points (median value was 7). Willet (1998) suggested that most investigators used a range of an FFQ between 5 and 10 and that five choices are likely to be too few and will usually result in a serious loss of information. Furthermore, as a rule of thumb, a scale of ten (never, once a month or less, 2-3 times per month, once per week, 2-4 times per week, 5-7 times per week, over one per day, 23 times per day, 4-6 times per day, over 6 times per day) for the FFQ seems reasonable because it provides detailed description on the high frequency end. Nevertheless, a food consumed less than once per week makes relatively little contribution to nutrient intake (Willet, 1998). Moreover, as mentioned in Chapter 2, PCA is a linear procedure which provides more accurate estimates when variables are approximately continuous and normally distributed, so a detailed scale of 10 items for the FFQ is more appropriate than smaller scales.

In observational studies in nutritional epidemiology it is advisable to adjust for energy intake (Willet 1998, Jakes et al, 2004), as this may itself be a risk factor, may alter the effect of a food or a nutrient on the potential outcome, or may reflect the variation of nutrient intake between individuals. In our systematic review, a few studies adjust for energy intake using the residual method as advised by Willett (1998) (Table B). However, Northstone et al. 2008 (Northstone, Emmett \& Rogers, 2008) demonstrated that it is not necessary to adjust for energy intake before entry into a PCA analysis to determine dietary patterns when using food frequency questionnaire data and that effects of total energy intake can be estimated at a later stage in the analytical process. In addition, in another study (DiBello et al., 2008) adjustment of the input variables for energy intake by the residual method seemed not to affect the factor
solution. This agrees with results from an older study (Balder et al., 2003), where the principal component solution was not substantially affected by energy adjustment.

## Decisions on how to derive a realistic number of dietary patterns with the use of PCA

The majority of the studies in our systematic review and in the review by Newby (2004) claimed that they derived a number of dietary patterns by
i) Examining the percentage of total variance that the dietary patterns explain in the original dataset. As mentioned in paragraph 3.2.7, the majority of the studies in our systematic review derived from 2 to 10 (median: 3, IQR: 2-4) dietary patterns which altogether explained a percentage of total variance with a median value of $24 \%$. The majority of these studies explain around $20 \%$ or $30 \%$ of the total variance (Table C), so the subset of principal components that is chosen contains little information about the variance of the original food intake variables. It is crucial to know how small percentage of total variance can be taken as serious information loss, and it is something that it needs further research in PCA studies of nutritional epidemiology.
ii) Employing the Kaiser criterion, that is retaining all the dietary patterns with corresponding eigenvalues above 1 . However, as mentioned in paragraph 2.6 in nutritional studies, a number of studies that decided the number of patterns based on the cut-off point for eigenvalues used a different cut-off point (median value:1.6, IQR:1.25-2) in order to retain a small and interpretable number of dietary patterns.
iii) A screeplot.

Our systematic review shows that the number of dietary patterns derived with the use of PCA are based on arbitrary notions of how meaningful and interpretable is a dietary pattern rather than to any of the other objective criteria mentioned above. Interpretation of the dietary pattern could be aided with the use of component loadings, component correlation coefficients and the appropriate method of rotation. However, again these methods rely on the subjective judgment of the researcher.

For example, Lutsey (2008) examined associations between dietary intake and metabolic syndrome after evaluation of the eigenvalues and their interpretability. All values $>2.0$ were retained, resulting in a 2 -factor solution and principal component loadings with absolute
values $\geq 0.20$ were shown in a table. If a lower cut-off point for eigenvalues was used, a larger number of dietary patterns would be derived. Furthermore labeling of the patterns would be different if a higher cut off point was selected for principal component loadings; "prudent" pattern would be characterized from fewer vegetables and not from yogurt.

Moreover, from our analysis in paragraph 6.3.2.1 if we strictly derived our dietary patterns according to the criteria mentioned above which are usually being applied in PCA studies then

- 129 dietary patterns should be derived if we wanted dietary patterns to explain a percentage of total variance of the original food items $>80 \%$, ( data are not shown)
- 69 dietary patterns should be derived if we wanted to retain all eigenvalues $>1$ (data not shown).
- 2, 6, 9 or 11 dietary patterns should be derived according to the screeplot

Label of our dietary patterns will also be different if

- We used a cut-off point of $>0.5$ or $<-0.5$ for our component correlations, our pattern could be easily consisted only of 3 fruits and 2 vegetables and labeled differently (see table 6.3.2.2.1)
- We used a different method of rotation than varimax; this had an effect on our results since slightly different patterns were derived. However, rotation of the components doesn't lead to different dietary patterns (Bountziouka et al. 2012)


## Standardisation of PCA to improve comparability between studies

Overall, our systematic review highlights the fact that it is the investigator's decision of how many dietary patterns should be derived and what food items are considered to be highly correlated with them in order to be interpretable. Furthermore, there are differences between the studies of how they are reporting empirically derived dietary patterns. This inherent methodological subjectivity of PCA and the lack of detailed reporting in every step of the application of the method in the literature is an important reason for observing inconsistencies between the studies. Improved reporting of PCA and additional
methodological studies on the method, as applied in nutritional epidemiology, will help to decrease the impact of subjectivity on the findings. Investigators must report analytically all decisions that were made in each stage of the analysis; beginning with the grouping and managing of the dietary data, with how patterns are presented and analyzed and how this affects the interpretation of our results and study findings. Some suggestions for directions of how the PCA method should be applied in nutritional epidemiology are summarized below with the use of Table E at the end of the thesis.

Food frequency questionnaires were used in $93.2 \%$ of the papers identified from our systematic review. Our view is that this is due to the fact that FFQ are easy to administrate and inexpensive survey instruments. However, FFQs are prone to measurement error and this need to be taken into account before or after the principal component analysis when using the instrument (Kipnis et al., 2008). Grouping food items into food groups is reported in the $28 \%$ of the studies. Our view is that by grouping several food items into food group's it i) increases the number of subjective decisions during the application of PCA especially since the multivariate statistical approaches are actually data reduction techniques, and ii) more importantly leads to information loss or underestimation of the results (McCann et al., 2001b). Although median number of the scale of the FFQ was 7, using more sensitive scales in the FFQ instrument could give a better description of diet in a continuous scale and a better approximation of population diet, since PCA was originally developed for the normal multivariate distribution and samples from it. A scale of 10 as a rule of thumb in the FFQ instrument is suggested in order variation of diet consumption of individuals to be captured (Willet et al, 1998). Median value of the items included in the FFQ instrument was 92. Our view is that food frequency questionnaire should comprise as many food items as possible with adequate contribution to nutrient intake (McCann et al., 2001b) to examine adequately the food consumption of the population under study. Our systematic review presented a small number of studies that converted food frequency data into grams/day (13.4\%) and fewer studies standardised their food intake data ( $4.2 \%$ ). The way that the food frequency data are converted and which food composition tables were used is essential for the dissemination of the results. In addition, since FFQ dietary data are highly skewed and could express different measurements of food item intake, (e.g. bread could be measured in grams/per day and beer in pints /per day), standardisation of the food intake variables is required and is strongly advised (also see paragraph 2.4.4). Very few studies reported adjustment for energy intake
(3\%) with residual method before the PCA. Our view is that because total energy intake could potential alter the results between the outcome and the exposure (Willet et al., 1998) should be included as a confounder in the analysis.

Scree plot provides an objective tool for deciding on the number of Principal components or dietary patterns to be derived. Other criteria's, such as the total percentage of variance being explained should be above $80 \%$ and eigenvalues should be above one, could result in a large number of dietary patterns which could be difficult to be interpreted. However, investigators should always report the total variance being explained by the dietary patterns, since as mentioned above a small percentage could be considered as serious information loss. Principal components were rotated in the majority of the studies (55\%). Rotation method doesn't seem to affect dramatically the derived dietary patterns; however different methods should be carried out in the studies and not only varimax for validity purposes. In addition, studies should always report if their method of rotation results into uncorrelated components or not (i.e. varimax rotation introduces correlation between the components to enhance interpretability). Using a cut-off point of 0.3 for principal component loading is not an objective tool for labeling the dietary patterns. This arbitrary cut-off point is derived from the fact that moderate and high correlations between variables are defined with a value higher than 0.3. Investigators should always examine lower or higher cut off points for labeling their dietary patterns. In addition, inconsistencies in the findings may be reduced if studies report dietary patterns quantitatively and not qualitatively.

Validation studies that include a reproducibility component are warranted for both methods examining the identified dietary patterns in a sub-sample of the original sample as well as using confirmatory factor analysis if there is a strong a-priori hypothesis could support the original findings of PCA. However more validation methods should be conducted in the studies (only $24 \%$ of the studies validated the reproducibility of their dietary patterns). Finally, additional research is needed on how decisions at the different stages of PCA could affect the dietary pattern solution.

### 7.3.3 Translation of posterior dietary patterns identified by PCA into an intervention

The relative importance of dietary patterns contributing to the effect of the overall dietary score cannot be ascertained without examining components included in the score. For example, the US Department of Agriculture has proposed the Health Eating Index (HEI) as
an index for monitoring dietary quality in the United States (Kant, 2004). However, a score based on several components makes it impossible to determine which areas of the diet need attention. Individual components contributing to the pattern must be examined to determine which ones should be targeted for intervention. PCA provides a similar score of a pattern based on the individual's consumption on a variety of foods and this score is associated with several health outcomes. However, it is unclear whether foods that are singled out from the PCA method and contribute more to the dietary pattern could make altogether a good dietary intervention.

### 7.3.4 Conclusions

Dietary patterns empirically derived with the use of PCA are difficult to replicate across studies because of a varying number, range, and type of variables estimated from a variety of dietary measurement methods before entered into the analysis procedure. Similarly, differences in the statistical analytical decisions and in the labelling of patterns across studies lead to incomparability as we observed in the previous sections. Qualitative and quantitative results from our systematic review present all the major shortcomings that PCA has in nutritional epidemiology, and provide strong evidence of support to our fundamental PhD assumption that less complicated methods than PCA could be equally or more effective at detecting diet-disease associations and provide better guidance for designing clinical trials and improve comparability of the findings between studies.

### 7.4 Interpretation of results from our simulation study

A simulation study was designed to evaluate the usefulness of PCA in dietary pattern analysis and PCA was compared to an ESFA procedure.

In some scenarios ESFA had greater power than PCA to detect an association of diet with disease. Allowing for multiple testing using a Benjamini and Hochberg approach, ESFA also typically had higher power and lower FDR for identifying the combination of foods that were causally linked with disease than a PCA in which combination of foods were singled out if they correlated highly ( $>0.3$ or $<-0.3$ ) with a significant dietary pattern. However, unadjusted ESFA and PCA had an uncontrolled false discovery rate, which increased with increasing power. This was the result of confounding between foods, many of which were highly correlated with each other. The false discovery rate of ESFA could be controlled at a fixed
level by adjusting for foods that were significant in an unadjusted ESFA, with surprisingly good power.

Even when a simplified "Western" dietary pattern was the real culprit, PCA could not outperform ESFA in reconstructing the foods that were linked with disease. Allowing for multiple testing, ESFA had greater power and similar false discovery rate with PCA for identifying combination of foods which constitute the simplified "Western" dietary pattern that was causally associated with disease. False discovery rates were extremely high for both methods. However, again ESFA could be controlled at a desirable fixed level of $20 \%$ by adjusting for foods that were significant in an unadjusted ESFA. These findings were replicated in two different FFQ data-sets.

Slattery et al. 1998 (Slattery \& Boucher, 1998) concluded that the soundness of using PCA to identify eating patterns would be better understood when more epidemiologists had begun to use the method and when a thorough understanding of the individual data elements has been obtained. However, there are no clear conclusions to be drawn on the usefulness of the method from research conducted during the last 13 years as Newby (2004) and our systematic review presented. One of the main reasons is the vague language used to justify PCA.

As mentioned above, one of the main justifications that researchers have used is that dietary patterns analysis with the use of PCA allows for the examination of the additive effects of dietary exposures which are too small to be detected by their own, or are diluted by confounding (Jacques \& Tucker, 2001, Randall et al., 1990). We would formalise these views in two ways - a "weak" sense and a "strong" sense.

The "weak" sense would be that PCA as a method is more likely than analysis of single foods to detect an additive effect of diet if there is one. As we observed in tables in paragraph 5.2 by estimating the power with which ESFA and PCA could detect any association between diet and disease, this was not the case. ESFA had better power in detecting any statistically significant effect between diet and disease, in $85 \%$ of our cases.

The "strong" sense would be that if disease is associated with a combination of intake of W and intake of $X$ and intake of $Y$ and intake of $Z$, then we are more likely to find effects of some or all of W, X, Y and Z by doing a PCA than by detecting this combinations of foods with the help of ESFA. To examine this, we estimated the power and false discovery rate
with which ESFA and PCA could detect specific combinations of foods which are causally linked to disease. As we can observe from the tables in paragraph 5.3, ESFA was more effective (higher power and lower or similar FDR) at detecting associations in this "strong" sense than PCA allowing for multiple testing and adjusting for energy intake in $93 \%$ of our cases.

It is common to try and control the FDR at a low level (Benjamini \& Hochberg, 1995). We have used a nominal FDR of $20 \%$; in genetic studies, where the use of FDR is wellestablished, FDRs between $5 \%$ and $20 \%$ are recommended depending on the circumstances. This corresponds to a search for candidate-gene effects requiring further replication, rather than for definitive evidence (Benjamini \& Yekutieli, 2005). Note that 20\% is still well below the FDR of individual hypothesis testing using $\mathrm{P}<0.05$ as a cut-off (Ioannidis, 2005). However, it is concerning that the observed FDRs of ESFA (nominally controlled at 20\%) and of PCA both increase in an uncontrolled fashion as the sample size and power increase (Tables in paragraph 5.3). It is worth noting that when one in seven foods are causally linked with disease, as here, an FDR of around $86 \%$ would be achieved by selecting "significant" foods entirely at random. The uncontrolled FDR occurs because all food intakes are correlated to some extent with the causal foods, leading to false positive findings (more so as power increases). More formally, the general conditions of Benjamini \& Hochberg procedure are not met, and therefore we cannot expect the FDR to be controlled at $20 \%$. In fact, as the sample size and power increases, it becomes easier and easier to detect all these indirect effects of other foods that are associated with the foods that have a causal effect, so we get more and more significant results, and more of these are false discoveries, because they are only indirect effects. So the FDR gets higher with increasing sample size, up to the maximum value it can take. We tried a variety of approaches to control for other foods in ESFA, and found that the FDR could be successfully controlled at the nominal level by adjusting for foods which were significant in a univariate analysis. To achieve given power, this required around twice the sample size of an unadjusted ESFA (with its inflated FDR), and four times the sample size from our original calculation, i.e. for an unadjusted analysis with criterion $P<0.05$. The problem with this adjustment is that with small sample sizes, the method won't always be able to make the correct adjustment. As the sample size increases we get more power to detect the things we should be adjusting for, so the adjustment becomes more effective. However, there are two competing effects: firstly, as sample size increases the FDR
tends to increase, because we get more power to detect indirect effects of foods; secondly, as sample size increases the adjustment becomes more effective, so the FDR starts to be controlled at the desired level. This is why we see FDRs increasing above $20 \%$ for lower sample sizes of 300,600 and 1200 and then decreasing again even below $20 \%$ for large sample sizes of 2400 and 4800 .

Our models included 4 or 10 or 30 foods or 2 different simplified "Western" dietary patterns (consisted of 30 foods from the FLAG study and 10 in the ECRHS II study) which were causally linked with disease. In each scenario, this corresponds to around 1 in 7 foods ( 30 foods in FLAG study and 10 foods in the ECRHS II study and 30 foods consisting of the simplified "Western" dietary pattern in FLAG study) and 1 in 20 foods ( 10 foods in FLAG study and 4 foods in the ECRHS II study and 10 foods consisting of the simplified "Western" dietary pattern in the ECRHS II study). Similar qualitatively findings were obtained from all of our simulations. We did not consider interactive effects of foods in our models: this requires further investigation. As mentioned in Chapter 2 dietary constituents interact with each other in complex ways to impact health, for example, on breast cancer risk and hypertension (Messina et al., 2001, Sacks et al., 1995). PCA is often recommended as a way of dealing with interactions between foods (Hu, 2002, Varraso et al., 2009). However, as was mentioned in paragraph 2.5, it is questionable whether the linear combinations of food intakes produced by PCA adequately address the issues of modelling interactions.

Although we considered a model based on a "Western" dietary pattern, there is no reason why food intakes with truly causal effects should be only food intakes that are highly correlated with each other. Hence, in our simulation study, foods may or may not be correlated between each other when they are associated with disease risk. Where disease risk is explained by other factors that are confounded with diet, this confounding is likely to be also at the level of a dietary pattern. As we highlighted in paragraph 7.3, a "prudent" dietary pattern, for example, is associated with older age (Agurs-Collins et al., 2009), female sex (Robinson et al., 2009), non-smoking (Fung et al., 2001), higher income (Perrin et al., 2005), higher educational level (Raberg Kjollesdal, Holmboe-Ottesen \& Wandel, 2010), exercise (Lopez-Garcia et al., 2004) and supplement use (Heidemann et al., 2008). These factors are likely to be associated with a number of food intakes contributing to a "prudent" pattern rather than with any one of these foods in particular. This is another reason to adjust each food effect for others found to be significant, as we suggest: this should help control for both
measured and unmeasured confounding. We did not include non-dietary confounders in our models because there are just too many different potential confounders and models for their effects that might be considered. We suspect, however, that as long as there are foods with truly causal effects, the findings presented in this paper will generalise to situations where other confounders have been explicitly adjusted for.

There is an asymmetry in the patterns of our Monte Carlo simulation results when different models (1, 2 and 3) have an effect on disease for ESFA method. Although the numbers of affected individuals are slightly lower for models where the power estimates are lower, this doesn't seem to be the only explanation for this asymmetry (see Tables 5.2.1, 5.2.2,). Models with all positive effects give different results to models with all negative effects; the situation that is being simulated is not symmetric, for example, the distribution of any given food intake cannot be expected to be symmetric but positively skewed as we observed in Figure 4.11.1. Thus, if the power to detect any effect of diet is much greater when all effects are positive than when all effects are negative, it is expected that when there is a mixture of positive and negative effects we could get intermediate power, because there are still a number of positive effects, which are easier to detect than negative ones. Similar patterns are observed when we are estimating the power and FDR of ESFA to detect the specific foods that are causally linked to disease (Tables 5.3.1 an 5.3.2).

However, average percentages of power of PCA in Tables 5.2.1, 5.2.2, 5.3.1 and 5.3.2 present a more symmetrical pattern when different models (1,2 and 3) have an effect on disease. This is due to the fact that principal component scores are not expected to be as positively skewed as food intake data (Figure 4.11.1). However, we could not find an adequate explanation why this is not consistent for the ECHRS II survey data-set simulation scenario in tables 5.2.2 and 5.3.2.

Furthermore, in tables 5.2.1, 5.2.2 and 5.2.3 power increases with the number of principal components and sometimes it decreases. For example, power of PCA is lower when results for 10 principal components are compared to results for 2 and 5 principal components. This is only reported when power of PCA outperforms ESFA for 2 and 5 principal components. When power of ESFA outperforms PCA for 2 and 5 components then power of PCA for 10 principal components increases. This pattern in our simulation results is observed because as
the number of principal components increase the power of PCA shifts towards the power of ESFA.

There is a growing interest in designing dietary interventions around foods rather than nutrients (Jacobs et al., 2009) and around particular foods rather than dietary patterns (Jacobs et al., 2009, Mann, 2010). Specifically, Jacobs suggests that the evidence for beneficial effects of a "prudent" diet comes from interventions which only modified the intake of one or two foods (Jacobs et al., 2009). Mc Cann (2001) found that fruits and vegetables alone provided the highest discrimination among endometrial cancer cases and controls compared with the other methods of characterisation and highlighted the fact that sophisticated techniques may be unnecessary in studies of diet and disease. Mann and Aune, evaluating the evidence that fruit and vegetables can prevent diabetes, have called for more studies looking at effects of specific fruits and vegetables (Mann, 2010).

In our study, we decided to vary the particular parameters that affect the decision of a researcher when employing PCA in nutritional data. Our main sources of information were our systematic review for years 2004-2009 (chapter 3 of the thesis) and two other reviews published previously (Kant, 2004, Newby \& Tucker, 2004). We decided that varying the number of our principal components from 2 to 10 components was reasonable since the upper limit for the derived principal component that were used by the literature was 10 . We decided to use an odds ratio of 1.5 , because this is a reasonable effect in epidemiological studies which are using dietary patterns and a false discovery rate of $20 \%$ in order not to penalize too heavily our findings.

Since ESFA outperforms PCA in our simulation study, dealing with high-dimensional multivariate dietary exposures could be treated as a problem of variable model selection that is, finding the nonzero regression coefficients in an unknown regression model. Our adjusted ESFA is similar to the iterative sure independent screening method (ISIS) for ultra-high dimensional data (Fan \& Lv, 2008). Other forms of penalised likelihood estimation methods have been developed in the last decade to cope with high-dimensional data and have been lately reviewed by Fan \& Lv (2010). This could be potentially useful in nutritional epidemiological studies, and further research is needed.

In conclusion, an FFQ-wide study of associations between food intakes and disease risk out-performs an analysis of dietary patterns derived from PCA. Analysing each food
adjusting for others allows truly causal effects to be identified with a low rate of false discoveries, and surprisingly good power. Although PCA has proved extremely popular in nutritional epidemiology to date, our simulation study questions its routine use in this context.

### 7.5 Analysis of $\mathbf{G A}^{\mathbf{2}} \mathbf{L E N}$ dataset: effect of diet on asthma

In this multicentre cross-sectional study, empirically derived dietary patterns ("fish, fruit and vegetables" and "meat, potatoes and sweets") with the use of a meta-analytic approach to PCA were not associated with respiratory symptoms, after controlling for potential confounders. In our two-step exhaustive single food analysis approach, several food items were statistically associated with respiratory and allergic symptoms. However, our two-step ESFA results were not all of them in line with biological plausible hypotheses.

Our non-statistically significant findings of associations between our dietary patterns and asthma were consistent with other observational studies of 54,672 French women (Varraso et al., 2009), 52,325 male and female adult Chinese Singaporeans (Butler et al., 2006) and 1453 adults living in Greenwich (Bakolis et al., 2010). However, in two large US prospective cohort studies of 42,917 men and 72,043 US women, weak associations were reported between a "prudent" pattern and adult onset asthma for the female population only. There were no associations with the other dietary patterns and adult onset asthma (Varraso et al., 2007a, Varraso et al., 2007b). Furthermore our results were consistent with the results of 12008 pregnant women in the Avon Longitudinal Study of Parents and Children (ALSPAC) (Shaheen et al. 2009), where no associations were observed between a "health conscious" (similar to our "fruit, fish and vegetables") and "processed" (similar to our "meat, potatoes and sweets") pattern and asthma, atopy and eczema. Comparison with other studies of adults and respiratory and atopic outcomes is difficult as no other studies have analysed dietary patterns using PCA in this setting.

The heterogeneous effect on the association of the "fruit, fish and vegetables" pattern and eczema and allergic rhinitis is not easy to explain. Diet is strongly socially, environmentally and culturally patterned with specificities between countries and social groups (Galobardes, Morabia \& Bernstein, 2001, Varasso, 2012). As Shaheen (2009) highlighted in the results from the Avon Longitudinal Study of Parents and Children (ALSPAC), a "health conscious" diet was univariately associated with eczema, total $\operatorname{IgE}, \mathrm{FEV}_{1}$ and negatively associated with persistent wheeze and asthma. However on controlling for numerous potential confounders these effects were attenuated and become non significant; went from an odds ratio of 0.90 ( $95 \%$ CI: $0.84-0.96$; p-value $<0.001$ ) to an odds ratio of 0.96 ( $95 \% \mathrm{CI}: 0.88-1.05$; p-value
$=0.37$ ) for the "health conscious" pattern; and from an odds ratio of 1.14 ( $95 \% \mathrm{CI}: 1.07-$ 1.22; p-value $<0.001$ ) to an odds ratio of 1.02 ( $95 \%$ CI: $0.94-1.10$; p-value $=0.69$ ) for the "processed" one. Heterogeneity in multi-centre studies can also suggest alternative explanations for apparent effects of diet observed in single centers, such as uncontrolled confounding, and would make us cautious of progressing to a trial (Burney et al., 2008).

Because a dietary pattern is acting as a proxy to individual foods associated with respiratory outcomes then some heterogeneity in its effect might be due to heterogeneity in its associations with these foods; there was a great amount of variation in the correlations of individual foods with the "fruit, fish and vegetables" pattern as observed in table 6.3.2.2.1.

A meta-analytic approach to deriving dietary patterns across a number of centres has been investigated once before (Hooper et al., 2010). This method can be successful in identifying common dietary patterns, as well as evidence for heterogeneity in the effects of those patterns. Heterogeneity in observational studies of diet can sometimes argue against progressing to trials.

When ESFA was applied to our data, we found protective effects of fruits, whole grains, nuts, fish and vegetables on asthma (any legumes, chard, cherries, radish, lemon, wholemeal bread), atopy (parsnip), chronic sinusitis (apple, fenugreek, cabbage, fresh white fish, raisins) allergic rhinitis (peach, rhubarb, smoked fatty fish, asparagus, kiwi, peanuts) and eczema (bitter melon, rhubarb). Hard fruit intake has been negatively associated with asthma (Shaheen et al., 2001) impaired lung function (Butland, Fehily \& Elwood, 2000) and chronic obstructive pulmonary disease (COPD) (Tabak et al., 2001a, Shaheen et al., 2010). However, longitudinal evidence for hard fruit intake relate only to COPD (McKeever \& Britton, 2004). In addition, vegetable intake (carrots, tomatoes, leafy vegetables) has been related to reduced asthma (Romieu et al., 2006). There is evidence that intake of fish may protect against asthma symptoms (Laerum et al., 2007) and impaired lung function and COPD in adults (Tabak et al., 2001b). However, results on the beneficial effect of fish intake have been contradictory (Thien et al., 1996). Finally, fish, fruits and vegetables are essential components of a Mediterranean diet, which other recent work has found to be associated with improved asthma control in adults (Barros et al., 2008).

These protective associations are in line with current hypotheses and may be explained by the high antioxidant content of these food items (Halliwell, 1996). High levels of vitamin C
and flavonoids are present in citrus and hard fruits, phenolic acids, flavonoids, phytic acid, avenanthramides, vitamin E and selenium in whole grains (Devereux \& Seaton, 2005, Slavin, 2004, Seaton, Godden \& Brown, 1994), and n-3 fatty acids, n-6 fatty acids in fish (Schnappinger et al., 2009, Sausenthaler et al., 2009). However, these results are not confirmed by experimental studies where no effect has been demonstrated for Vitamin C, E and magnesium on asthma (Pearson et al., 2004, Fogarty et al., 2003).

Contrary to prevailing paradigms, we also found some evidence, for a positive association between fruits and vegetables intake with respiratory and allergic outcomes and a negative association between condensed milk with asthma and eczema. Recent studies suggested that a higher intake of antioxidants might promote the development of allergic disease (Murr et al., 2005) and whole milk intake might reduce asthma symptoms (Woods et al., 2003). However these results need further replication.

The explanation of other associations found between food items and respiratory and allergic outcomes are hard to be explained. Exhaustive single food analysis identified a number of foods which have not previously been considered to be associated with potential respiratory outcomes, and there are no obvious mechanisms that might link them. Although some of these may represent new and important findings, our analysis is limited by the presence of unmeasured, reverse and qualitative confounding.

Specifically, a lot of the highly significant foods have odds ratios close to 1 , so it could be argued that the smallest bit of unmeasured or residual confounding or bias might explain their associations with respiratory and allergic outcomes. A simulation study by Fewel \& Davey Smith \& Sterne (2007), showed that bias in the exposure effect estimate increases as the amount of residual and unmeasured confounding increases, especially when confounders are uncorrelated with each other. Adjusting only for age, gender, body mass index and smoking status and food intakes that were statistically significant in an unadjusted ESFA didn't seem to control adequately for potential confounders.

In addition, we did not control our analysis for total energy intake. As we highlighted in paragraph 4.3.1, adjustment for total energy intake is important and should be considered because the level of energy intake might be a risk factor and distort the effect of a food on the potential respiratory and allergic outcome (Willett et al., 1997). Furthermore the effect of diet on allergic disease should be adjusted for other factors related to individual characteristics
such as social class, level of education, culture, ethnic background, language, as well as psychological factors (Tricon et al., 2006). In a systematic review by Nurmatov et al. (2012) on the effect of confounding in studies of diet and allergies in children, authors suggested that future studies should be adjusted for maternal characteristics, birth measurements, socioeconomic characteristics, other dietary factors and environmental exposures. Unfortunately in our analysis, information on cultural, socioeconomic and psychological factors was not available at the time of the analysis.

One other limitation is that there are differences in the correlation matrices between the centres. It is very difficult to present correlations between 238 food items for each of the ten centres. However, we presented in the thesis correlations of individual food items with our meta-analysed dietary patterns (Table 6.3.2.2.1) empirically derived with PCA. This allows us to investigate whether the correlation structure looks the same in different centres. Correlations of each food item with the dietary pattern are very different in each centre and this has implications in our meta-analysis results. In each centre, the higher the difference in the correlation values between food items, the higher the difference in the correlation matrices that are used in the PCA and the greater the heterogeneity of the dietary patterns that are meta-analysed; thus, it is doubtful the appropriateness of pulling dietary patterns from all the centres together. Appropriate statistical methods to derive and synthesize dietary patterns among different centres is a topic for further investigation.

Furthermore, because of the cross-sectional nature of our study there is the possibility of reverse confounding; people with respiratory symptoms may alter their diet in order to be healthier. For example, this is highlighted when we stratified our analysis for non atopics, associations between any legumes, chard, cherry, lemon, wholemeal bread and asthma didn't remain statistical significant. Finally, qualitative confounding is also apparent to our analysis; specifically in the univariate analysis chard is assumed to increase the risk of asthma, but when we adjust for the food intakes that were statically significant in an unadjusted ESFA (analysis not adjusted for other foods, but adjusted for age, sex, body mass index and smoking status) chard has a protective effect on asthma (Table 6.3.4.1).

FDR is controlled at $20 \%$ and $5 \%$, so these are all foods for which there is some evidence of an association with health outcomes, not a clear definitive list. In our analysis, since we control the rate of our false discoveries at $20 \%$ and $5 \%, 20 \%$ and $5 \%$ of those discoveries are
expected to be false and consequently $80 \%$ and $95 \%$ are expected to be true. So, for example, in the results of the adjusted ESFA after controlling our FDR at 20\%, 18 foods were expected to be associated with asthma (Table 6.3.4.1). Hence we can expect 14 of these to be genuinely associated with asthma and 4 of them to be false discoveries. Similarly, we can expect 11 (out of 14) foods to be genuinely associated with chronic sinusitis, 12 (out of 15) with allergic rhinitis, 8 (out of 10 ) with eczema and 14 (out of 17) with atopy. As expected, we had a smaller list of food items when we control our FDRs at $5 \%$. Majority of these foods were indentified when FDR was at $20 \%$. However, again a lot of these associations couldn't be easily clinically explained; positive association between couscous and turnip with asthma; positive associations of okra and pumpkin with chronic sinusitis; negative association of moussaka with atopy.

In order to examine associations between diet and disease, we employed a random intercept logistic model to take into account the dependence of individuals between countries. One of the major assumptions of the tests of significance used in the multilevel models is normality of the error distributions involve. Another assumption is the sufficient number of sample size for the higher-level variables. A potential limitation for using the multi-level model approach is that our countries are not a random sample of a population of countries, and that the number of our countries in the sample may be small. However, random intercept and random coefficient models are a flexible and powerful way to tackle the effect of clustering in country-level data (Localio et al., 2001) and it has been suggested that 10 or more groups in the second level (in our case country-level), provide accurate standard error estimates for the regression coefficients of the fixed part of the multilevel model (Maas and Hox, 2004).

An additional analysis was performed for the investigation of the consistency of effects across countries, reported in our adjusted ESFA controlling the FDR at 5\%, of food items on respiratory and allergic outcomes. Hence, two types of approaches were employed: a random effects model and a meta-regression approach. From the random effects model all associations remained statistically significant with similar effect sizes (see Appendix IX). In the meta-regression model only the association between turnip and asthma appeared to be heterogeneous between centres ( $\mathrm{I}^{2}=57.4 \% ; \mathrm{P}<0.05$ ).

Comparing the results of the meta-regression approach with the results from the random intercept and the random effects modeling approach we observed that effect sizes were
similar. Associations that remained borderline significant was between; smoked fatty fish and asthma, couscous and asthma (see Appendix III); okra and chronic sinusitis (see Appendix IV); peach and allergic rhinitis, vegetable oil and allergic rhinitis, smoked poultry and allergic rhinitis, sour cream and allergic rhinitis (see Appendix V); rhubarb and eczema, crisp fried cakes and eczema, bitter melon and eczema, Greek style yogurt and eczema (see Appendix VI); crisp fried cakes and atopy and moussaka and atopy (see appendix VII). However, multilevel and meta-regression techniques are conceptually different. Multilevel analysis fits a hierarchical model and estimates the effect of a food item on disease by using individual-level observations which are clustered within the countries, while meta-regression estimates a separate effect of a food item on disease in each country, and estimates an overall pooled effect based on these per country effect sizes. In addition, for our meta-regression approach associations reported are pretty consistent in all counties. This is an additional argument of these associations to be real findings and even close to be causal according to Bradford-Hill (1965) criterion of consistency (although a confounder might be closely attached to the same food in all countries).

Although the use of adjusted ESFA obtained a set of foods predictive of the respiratory status of the individuals, the foods identified failed to represent a number of foods which are in line with current biological hypothesis or present some new streams in the nutritional research. It's possible that confounding between lifestyle factors and individual foods explains some of the findings, and that some of the statistical methods of analysis are just not appropriate (such as the assumption of random effect of site in our regression models or the correlation structure of food intakes which is very different in different sites). Although, less complicated methods than PCA such as variable selection methods seem promising in the field of nutritional epidemiology, our G.A.L.E.N analysis results suggest that challenges remain in this field.

In conclusion, we found no firm, consistent evidence for an association of dietary patterns empirically derived with the use of PCA with respiratory and allergic outcomes. When we employed ESFA in our data-set, a number of foods were associated with respiratory and allergic symptoms. However, our results may be affected by unmeasured confounders associated with dietary choices and statistical methodological issues so they must be interpreted with caution and need further replication.

### 7.6 Conclusions

Slattery in 1998 concluded that the soundness of using principal component analysis to identify eating patterns will be better understood when more epidemiologists have begun to use the method. In 2008, ten years after the explosion of the use of PCA in dietary studies, and with almost 100 papers employing the PCA method, Slattery claimed that data-driven dietary patterns identified empirically from PCA characterized the diet associated disease risk, provided more seemingly consistent associations and could guide diet and disease investigations better than anyone food or nutrient. All of these arguments are questionable.

Specifically, our systematic review provides evidence of highly inconsistent data-driven dietary patterns, empirically derived with the use of PCA, across different studies. This may be due to the fact that dietary patterns are strongly confounded by lifestyle factors. In addition, reviewing the literature from our systematic review and the review from Kant (2004) and Newby (2004) we concluded that there is a number of different subjective decisions taken by the researchers (because of the methodological subjectivity of PCA as employed in nutritional epidemiology) from study to study. Hence, data-driven dietary patterns cannot be easily translated into an intervention and be comparable across different studies.

Our simulation study provides quantitative evidence that questions the use of PCA for detecting combination of foods that are causally associated with disease and highlights the high risk of false positive findings of the method. Moreover, we suggest that the best way to analyse nutritional data assuming that a combination of foods are associated with disease in an additive way or as a principal component of the population is to 1 ) run an exhaustive search for associations between individual food intakes and disease (ESFA procedure) with an appropriate statistical model, allowing for multiple testing and adjusting for energy intake and then 2) to re-run an ESFA procedure by further adjusting for these foods that are statistically significant in the first round of univariate analysis.

In the $G A^{2}$ LEN survey analysis study, PCR method was unsuccessful on detecting associations between diet and the allergic and respiratory outcomes. On the other hand, ESFA detected a number of associations, but not all of them were necessarily causal. Specifically, ESFA gives a list of variables with a $20 \%$ false discovery rate and not a definite list.

However, lack of biological plausible associations could be also due to unmeasured confounding or that assumptions of the statistical methods might not be met.

Our fundamental assumption was that in order for PCA and its extension PCR to be useful methods in nutritional epidemiology, they should identify combinations of foods in the population that are causally linked to disease, be reproducible and can be translated into an intervention. None of the three are possible with the use of the PCA and PCR methods as we proved from our systematic review and our simulation study.

We are not to claiming that different elements of diet are associated with disease only in the way that we specified or that there are not complex biological interactions between nutrients and foods or that all other data reduction techniques that associate a pattern with disease are not useful. The purpose of the thesis is to provide a critique and start a discussion on the inappropriateness of PCA as a prevailing method in nutritional epidemiology to detect associations between diet and disease, and emphasize on the need for future research (paragraph 7.7) in this field.

Paraphrasing Lehman on the theatre of Bertolt Brecht "We are highly interested in the questions that led to the explosion of dietary pattern analysis with the use of PCA, but we are not satisfied anymore by the answers provided by the method".

### 7.7 Ideas for future research and implications of these findings

Analysis of nutritional data is exceedingly challenging, given that the aim is to tease apart complex data patterns, estimate the effect of overall diet, and tackle confounding of dietary exposures and estimate possible higher-order interactions between nutrients or foods consumed in combination. Possible ways to address all these issues in the future without the use of a PCA approach are

- To evaluate other existing multivariate methods applied in nutritional epidemiology. This may be done by comparing Cluster analysis, Treelet transform, Reduced Rank Regression, Gaussian mixture modelling with an exhaustive analysis of single foods (ESFA).
- To expand our simulation model in a way that different foods are causally linked with disease in an interactive way, and evaluate existing multivariate methods by comparing them with an exhaustive analysis of single foods (ESFA).
- To use other datasets for our Monte Carlo simulations with people living outside UK.
- To research into new methods for analyzing nutritional data with the use of statistical techniques applied for high dimensional data such as Sparse PCA (Jolliffe 2003), supervised Principal Components (Bair et al., 2004), Iterative Sure Independence Screening (Fan and Lv, 2008), Shrinkage and Selection via the Lasso (Tibshirani, 1996 ) and causal models (Galea et al.2010, Vineis., 2006).
- To continue the analysis of G.A ${ }^{2}$.L.E.N data-sets for nutrients and other respiratory and allergic outcomes.
- To aggregate food items into food groups of the similar composition and exchangeability (e.g. biscuits and crackers) that are consumed each with an average frequency or less than once per week in G.A ${ }^{2}$.L.E.N.
- To apply our exhaustive single food analysis method adjusted for foods that were significant in an unadjusted analysis (analysis not adjusted for other foods, but adjusted for age, sex, body mass index and smoking status) in other observational studies apart from G.A ${ }^{2}$.L.E.N.


## Tables of systematic review

Table A. Justifications for the use of Principal Component Analysis instead of a single nutrient/ food approach in dietary pattern analysis studies in nutritional epidemiology (Sorted by Year of Publication and Author). Blank cells describe not available information on that argument. Proportions of papers that provide this justification is given.

| Author | Interactive, antagonistic and synergistic effects of foods consumed in combination. (proportion of papers : $25 / 163=15.3 \%)$ | Additive effects of foods consumed in combination which are too small to detect when they are examined separately (proportion of papers: $14 / 163=8.5 \%$ ) | Confounding and mulitcollinearity from lifestyle factors and dietary exposures (proportion of papers: 52/163=31.9\%) | Multiple testing problems (proportion of papers: 7/163=4.2\%) | Public health recommendations (proportion of papers: $14 / 163=8.5 \%)$ | Complexity of diet (proportion of papers: $25 / 163=15.3 \%)$ | Better evaluation / representation of overall diet (proportion of papers: $30 / 163=18.4 \% \text { ) }$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Agurs-Collins et al., 2009) |  |  | X |  |  |  | X |
| (Akbaraly et al., 2009) |  |  |  |  |  |  |  |
| (Ambrosini et al., 2009) |  | X | X |  |  |  |  |
| (Bakolis et al., 2010) |  | X | X | X |  |  |  |
| (Bastos et al., 2010) |  |  | X |  |  |  |  |
| (Bertuccio et al., 2009) |  |  | X | X |  |  |  |
| (Brantsaeter et al., 2009) |  |  |  |  |  |  |  |
| (Cottet et al., 2009) |  |  |  |  | X |  |  |
| (Craig et al., 2010) |  |  | X |  |  |  |  |
| (Cutler et al., 2009) |  |  | X |  | X | X |  |
| (Deshmukh-Taskar et al., 2009) |  |  |  |  | X | X |  |
| (Erber et al., 2010) |  |  |  |  |  |  |  |
| (Hamer \& Mishra, 2010) | X | X |  |  |  |  |  |
| (He et al., 2009) |  |  | X |  |  |  |  |
| (Hooper et al., 2010) | X |  | X |  |  |  |  |
| (Hughes et al., 2009) |  |  |  |  |  |  |  |

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| (Imamura et al., 2009) |  | X |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Jackson et al., 2009) |  |  |  |  |  |  |
| (Kesse-Guyot et al., 2009) |  | X |  |  | X |  |
| (Kontogianni et al., 2009) | X | X |  |  |  | X |
| (Qi et al., 2009) |  |  |  |  |  |  |
| (Lutsey et al., 2009) |  |  |  |  |  |  |
| (Muller et al., 2009) |  |  |  |  |  | X |
| (Nettleton et al., 2009) |  |  |  |  |  |  |
| (Noel et al., 2009) |  |  |  |  |  |  |
| (Oddy et al., 2009) |  |  |  | X |  | X |
| (Paradis et al., 2009) | X |  |  |  |  | X |
| (Reedy et al., 2010) |  |  |  |  |  |  |
| (Rezazadeh, Rashidkhani \& Omidvar, 2010) |  |  |  |  |  | X |
| (Robinson et al., 2009) | X |  |  |  | X |  |
| (Shaheen et al., 2009) |  | X | X |  |  |  |
| (Touvier et al., 2009) |  |  |  |  |  |  |
| (Uusitalo et al., 2009) |  |  |  |  | X |  |
| (Vujkovic et al., 2009a) |  |  |  |  |  |  |
| (Vujkovic et al., 2009b) |  |  |  |  |  |  |
| (Wiles et al., 2009) |  |  |  |  |  |  |
| (Wu et al., 2009) |  |  |  |  | X |  |
| (Ambrosini et al., 2008b)( 2 papers) |  |  |  |  |  | X |
| (Ambrosini et al., 2008a) |  |  |  |  |  |  |
| (Arkkola et al., 2008) |  |  |  |  | X |  |
| (Borland et al., 2008) |  |  |  | X |  |  |
| (Butler et al., 2008) |  | X |  |  |  |  |
| (Campbell, Sloan \& Kreiger, 2008) |  | X |  |  |  |  |
| (Chang et al., 2008) | X | X |  |  |  | X |
| (Crozier et al., 2008) | X | X |  |  |  | X |
| (De Stefani et al., 2008a) |  |  |  |  |  |  |
| (De Stefani et al., 2008b) |  |  |  |  |  |  |

Tables of Systematic Review

| (D'Souza et al., 2008) | X |  | X |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Edefonti et al., 2008) |  |  | X | X |  |  |  |
| (Engeset et al., 2009) | X |  |  |  |  | X |  |
| (Esmaillzadeh \& Azadbakht, 2008a) | X |  | X |  |  |  | X |
| (Esmaillzadeh \& Azadbakht, 2008b) | X | X | X |  |  |  | X |
| (Feinstein et al., 2008) |  |  |  |  |  |  |  |
| (Flood et al., 2008) | X |  | X |  |  |  |  |
| (Heidemann et al., 2008) |  |  |  |  |  | X |  |
| (Hooper et al., 2008) | X |  | X |  |  |  |  |
| (Iqbal et al., 2008) | X |  |  |  | X |  |  |
| (Keskitalo et al., 2008) |  |  |  |  |  |  |  |
| (Kim et al., 2008) |  |  |  |  |  | X |  |
| (Knudsen et al., 2008) |  |  | X |  |  |  |  |
| (Kubo et al., 2008) | X |  |  |  |  |  |  |
| (Lau et al., 2008) | X |  |  |  |  | X |  |
| (Lutsey, Steffen \& Stevens, 2008) |  |  |  |  |  |  |  |
| (McNaughton et al., 2008) | X |  |  |  |  | X |  |
| (McNaughton et al., 2008) | X |  |  |  |  | X |  |
| (Murtaugh et al., 2008) |  |  |  |  |  |  |  |
| (Nanri et al., 2008) (2 papers) | X |  | X |  |  | X |  |
| (Nettleton et al., 2008b) | X |  |  |  |  |  |  |
| (Nettleton et al., 2008a) |  |  |  |  |  |  | X |
| (Northstone \& Emmett, 2008a) |  |  |  |  |  |  |  |
| (Northstone, Emmett \& Rogers, 2008a) |  |  |  |  |  |  | X |
| (Northstone, Emmett \& Rogers, 2008b) |  |  | X |  |  |  | X |
| (Northstone \& Emmett, 2008b) |  |  |  |  |  |  |  |
| (Okubo et al., 2008) | X | X |  |  |  |  |  |
| (Romaguera et al., 2008) |  |  |  |  |  |  |  |
| (Sadakane et al., 2008) | X | X |  |  |  |  |  |
| (Shi et al., 2008) | X |  |  |  |  |  |  |
| (Takaoka \& Norback, 2008) |  |  |  |  |  |  |  |

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| (Tseng et al., 2008) |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Varraso et al., 2007b) | X |  |  |  |  |  |  |
| (Yannakoulia et al., 2008) |  |  |  |  |  |  | X |
| (Akesson et al., 2007) |  |  |  |  |  |  | X |
| (Bamia et al., 2007) | X |  | X |  |  |  |  |
| (Cai et al., 2007) | X |  | X |  |  | X |  |
| (Cui et al., 2007) |  | X |  |  |  |  |  |
| (Custodio das Dores et al., 2007) |  |  |  |  |  |  |  |
| (Dalvi, Canchola \& Horn-Ross, 2007) | X | X |  |  |  |  |  |
| (De Stefani et al., 2007) |  |  | X |  |  |  | X |
| (Esmaillzadeh et al., 2007) |  |  |  |  |  |  | X |
| (Hirose et al., 2007) |  |  |  |  |  | X |  |
| (Kim et al., 2007) |  |  |  |  |  |  |  |
| (Maruapula \& Chapman-Novakofski, 2007) |  |  |  |  |  |  |  |
| (Masala et al., 2007) | . |  | X | X |  |  | X |
| (McNaughton et al., 2007) |  |  |  |  |  |  |  |
| (Murtaugh et al., 2007) |  |  |  |  | X |  |  |
| (Nettleton et al., 2007) | X |  |  |  |  |  |  |
| (Okubo et al., 2007) | X | X |  |  |  |  |  |
| (Panagiotakos et al., 2007a) | X |  | X |  | X |  | X |
| (Panagiotakos et al., 2007b) |  |  |  |  |  | X | X |
| (Robinson et al., 2007) |  |  |  |  |  |  |  |
| (Sant et al., 2007) |  |  |  |  |  |  |  |
| (Shimazu et al., 2007) | X |  | X |  |  |  |  |
| (Shin, Oh \& Park, 2007) |  |  |  |  |  |  |  |
| (Takata et al., 2007) |  |  | X |  |  |  |  |
| (Teucher et al., 2007) | X |  |  |  |  | X | X |
| (Varraso et al., 2007a) |  |  |  |  |  |  |  |
| (Burt et al., 2006) |  |  |  |  |  |  |  |
| (Butler et al., 2006) | X |  | X |  |  |  |  |
| (Chen et al., 2006) |  |  |  |  |  |  |  |

Tables of Systematic Review

| (Crozier et al., 2006) | X |  | X |  | X |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Cuco et al., 2006) |  |  | X |  |  |  | X |
| (Kesse, Clavel-Chapelon \& Boutron-Ruault, 2006) |  |  | X |  |  |  | X |
| (Mizoue et al., 2006) | X |  | X |  |  |  |  |
| (Naska et al., 2006) | X |  |  |  | X |  |  |
| (Newby et al., 2006ba) |  | X | X |  |  |  | X |
| (Newby et al., 2006b) | X | X | X |  |  |  |  |
| (Pala et al., 2006) | . |  |  | X |  | X |  |
| (Paradis, Perusse \& Vohl, 2006) | X |  |  |  |  | X |  |
| (Ronco et al., 2006) |  |  |  |  |  |  |  |
| (Schulze et al., 2006) |  |  |  |  |  |  |  |
| (Waijers et al., 2006) | X | X | X |  | X | X |  |
| (Weismayer, Anderson \& Wolk, 2006) |  |  |  |  |  |  |  |
| (Wu et al., 2006) | X |  |  |  |  |  |  |
| (Zhang et al., 2006) |  |  |  |  |  |  |  |
| (Balder, Goldbohm \& van den Brandt, 2005) |  |  |  |  |  |  | X |
| (Cottet et al., 2005) | X |  | X |  |  |  |  |
| (Engeset et al., 2005) |  |  |  |  |  | X |  |
| (Fung et al., 2005) | X |  | X |  |  |  |  |
| (Hoffmann et al., 2005) |  |  | X |  |  |  |  |
| (Kroenke et al., 2005) |  |  |  |  |  |  |  |
| (Mannisto et al., 2005) |  |  | X |  | X |  |  |
| (Marchioni et al., 2005) | X |  | X |  |  | X |  |
| (Michaud et al., 2005) | X |  |  |  |  |  | X |
| (Mikkila et al., 2005) | X | X |  |  |  |  |  |
| (Mizoue et al., 2005) | X |  | X |  |  |  |  |
| (Montonen et al., 2005) | X |  | X |  |  |  | X |
| (Nkondjock et al., 2005) | X |  |  |  |  |  | X |
| (Northstone \& Emmett, 2005) |  |  |  |  |  |  |  |
| (Park et al., 2005) | X |  | X |  |  |  |  |
| (Perrin et al., 2005) | X |  |  |  |  | X |  |

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| (Rashidkhani et al., 2005) | X |  | X |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Uusitalo et al., 2005) | X |  | X |  |  |  |  |
| (Velie et al., 2005) |  |  |  |  |  |  |  |
| (Yang, Kerver \& Song, 2005) | X |  |  |  |  |  |  |
| (Corrao et al., 2004) |  |  |  |  |  |  |  |
| (De Stefani et al., 2004) |  |  |  |  |  |  |  |
| (Dixon et al., 2004) | X |  |  |  |  |  |  |
| (Fung et al., 2004a) |  |  |  |  |  |  |  |
| (Fung et al., 2004b) | X |  |  |  |  |  |  |
| (Kant et al. 2004) |  |  |  |  |  |  |  |
| (Khani et al., 2004) | X |  | X | X |  |  |  |
| (Kim et al., 2004) | X |  |  |  |  | X |  |
| (Lopez-Garcia et al., 2004) |  |  |  |  |  |  |  |
| (Newby et al., 2004) (Newby, Muller \& Tucker, 2004) (2 papers) |  |  |  |  |  |  |  |
| (Robinson et al., 2004) |  |  |  |  |  |  |  |
| (Sieri et al., 2004) |  |  |  |  |  |  | X |
| (Togo et al., 2004) | X |  |  |  | X |  |  |
| (Tseng et al., 2004) |  | X |  |  | X |  |  |
| (Wu et al., 2004a) |  |  |  |  |  |  |  |

Table B. Application of PCA procedure to identify dietary patterns in nutritional epidemiology; details of each study; details on the dietary assessment instrument being used in each study; preparation of data before entering the PCA procedure in each study; applications of criteria of PCA for labelling and identifying the number of dietary patterns in each study; validation methods for retained dietary patterns in each study (Sorted alphabetically by study design and country). Summary statistics are provided for each paper.


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|  |  | $\quad$ (median value (median value:38 ; IQR:19-69) 4. $\quad$ Scale of FFQ (median value: 7 ; IQR: $\quad \mathbf{5 - 1 0}$ ) |  | each study <br> 3. Scree Plot (proportion of papers: 81/163=49.1\%) <br> 4. cut-off point $s$ for eigen values (median value:1.6; IQR (1.2) <br> 5. dietary patterns interpretability (proportion of papers:70/163=42.9\%) <br> 6. Van der Voet's test (proportion of papers: $1 / 163=0.6 \%$ ) | (proportion of papers: 8/163=4.9\%) <br> 8. $\varphi$ coefficient for testing <br> inter-correlation <br> (proportion of papers: 2/163=1.2\%) <br> 9. Stricter cut-off points for energy intake (proportion of papers: $\mathbf{1 / 1 6 3 = 0 . 6 \%}$ ) <br> 10. Different method of rotation (proportion of papers: 2/163=1.2\%) <br> 11. Pearson Correlation coefficient in different time points (proportion of papers: 5/163=3.0\%) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Iqbal et al., 2008) | 1. INTERHEART study <br> 2. Case-control study <br> 3. 5761 Cases and 10646 controls <br> 4. 52 Countries | $\begin{aligned} & \hline 1 \mathrm{a} . \\ & 2.19 \end{aligned}$ | 1. | $\begin{aligned} & \hline 1 .>\|0.25\| \& 3 . \& 4 \\ & \text { Eigenvalue }>1 \& 5 \end{aligned}$ | , |
| (Ambrosini et al., 2008b) <br> (Ambrosini et al., 2008a) | 1. * <br> 2. Case-control study <br> 3. 546 Cases and 447 controls <br> 4. Australia | 1a. <br> 2.101 <br> 3.74 <br> 4.Scale of 10:'never" <br> to <br> " 3 or more times per day.' | * | 1.>0.3 2a \& 3. \& 4. >1. $\& 5$ | * |
| (Marchioni et al., 2005) | $\begin{array}{ll}\text { 1. } & * \\ \text { 2. } & \text { Case-control study }\end{array}$ <br> 3. 517 <br> 4. Brazil | $\begin{aligned} & 1 \mathrm{a} . \\ & 2.27 \end{aligned}$ | * | $1>02 \mathrm{a} .3 . \& 4 \& 5$. | * |
| (Campbell, Sloan \& Kreiger, 2008) | 1. National Enhanced Cancer Surveillance System <br> 2. Case-control study <br> 3. 2813 <br> 4. Canada | 1a. <br> 2.69 <br> 4. Scale of 10: from "never or less than once per month," to "six or more times per day" | * | 2a.. \& 3. \& $4>3 . \& 5$ | * |
| (D'Souza et al., 2008) | $\begin{array}{ll}\text { 1. } & \text { * } \\ \text { 2. } & \text { Case-control study }\end{array}$ | $\begin{aligned} & \hline 1 \mathrm{a} . \\ & 2.151 \end{aligned}$ | * | 1. >0.30 2a. \& $3 \& 4 \&$ 5 | * |

Tables of Systematic Review

\begin{tabular}{|c|c|c|c|c|c|c|}
\hline \& 3.
4. \& \begin{tabular}{l}
A total of 149 Cases and 251 controls were included for the study. \\
Canada
\end{tabular} \& 4. Scale of 8 "never or less than once per month" to " 6 per day." \& \& \& \\
\hline (Nkondjock et al., 2005) \& 1.
2.
3.
4. \& NECSS study Case-control study 585 Canada \& \[
\begin{aligned}
\& \hline 1 \mathrm{a} . \\
\& 2.69
\end{aligned}
\] \& * \& \[
\begin{aligned}
\& \text { 1. }>|0.30| \& 2 \mathrm{a} 3 . \& 4 \\
\& >1 \& 5
\end{aligned}
\] \& * \\
\hline (Cui et al., 2007) \& \[
\begin{aligned}
\& 1 . \\
\& 2 . \\
\& 3 . \\
\& 4 .
\end{aligned}
\] \& Shanghai Breast Cancer Study Case-control study 1556 controls China \& \[
\begin{aligned}
\& 1 \mathrm{a} . \\
\& 2.76
\end{aligned}
\] \& * \& \[
\begin{aligned}
\& \text { 1. }>|0.20| \& 2 \mathrm{a} 3 . \& 4 \\
\& >1 \& 5 \& 7
\end{aligned}
\] \& * \\
\hline (Di bello et al. 2008) \& \[
\begin{aligned}
\& 1 . \\
\& 2 . \\
\& 3 . \\
\& 4 .
\end{aligned}
\] \& Case-control study 3574 Cases and controls Costa Rica \& \[
\begin{aligned}
\& \hline \text { 1a. } \\
\& 2.135 \\
\& 3.43
\end{aligned}
\] \& 3 \& 5. \& 6 . \& 1. \\
\hline (Hooper et al., 2010) \& \[
\begin{aligned}
\& 1 . \\
\& 2 . \\
\& 3 . \\
\& 4 .
\end{aligned}
\] \& \begin{tabular}{l}
UK European Community Respiratory Health Survey I ( ECRHS-I) \\
Case- control study \\
1174 \\
Europe
\end{tabular} \& \begin{tabular}{l}
1a. \\
2.158 (German) \\
198 (UK) \\
204 (Norway174) \\
4.frequencies from never to five portions a day
\end{tabular} \& 1. \& 2 . \& 1. \(>|0.30| 2 \mathrm{a} . \& 3\). \& * \\
\hline (Bertuccio et al., 2009) \& 2.
3.
4.
4. \& Case-control study 230 patients with incident, histological confirmed gastric cancer and 547 frequencymatched controls, Italy \& \[
\begin{aligned}
\& \hline 1 \mathrm{a} . \\
\& 2.78
\end{aligned}
\] \& 1. \& \(1 . \geq|0.63| \& 2 \mathrm{a}\). \& 2. \& 4. \& 5. \\
\hline (Edefonti et al., 2008) \& 1.
2.
3.

4. \& Case-control study Cases were 2,569 breast cancers and 1,031 ovarian. Controls were 3,413 women from the same hospital network Italy \& $$
\begin{aligned}
& 1 \mathrm{a} . \\
& 2.78
\end{aligned}
$$ \& * \& \[

$$
\begin{aligned}
& \text { 1. }>0.63 \text { 2a \& } 3 . \& 4 . \\
& >1 . \& 5 .
\end{aligned}
$$
\] \& 7. <br>

\hline (Corrao et al., 2004) \& 1. \& Case-control study 481 selected controls, 152 were healthy subjects and 329 not Italy \& | $1 \mathrm{a}$ $2.93$ |
| :--- |
| 4.Scale of 10:(3 times/d, to never/rarely) | \& 3 \& 4.>1 \& 3. <br>

\hline (Jackson et al., 2009) \& 3. \& Case-control study 204 histological confirmed newly diagnosed prostate cancer Cases \& $$
\begin{aligned}
& \hline 1 \mathrm{a} . \\
& 3.33
\end{aligned}
$$ \& 1. \& \[

$$
\begin{aligned}
& \text { 1. }>|0.40| \& 2 \mathrm{a} . \& 3 . \& \\
& 4 .>1
\end{aligned}
$$
\] \& * <br>

\hline
\end{tabular}

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|  | and 204 individually matched urology clinic controls <br> 4. Jamaica |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Hirose et al., 2007) | 1. Aichi Cancer Center (HERPACC). <br> 2. Case-control study <br> 3. 1885 <br> 4. Japan | $\begin{aligned} & \hline \text { 1a. } \\ & 2.31 \end{aligned}$ | * | 2a 3. \& 4>1 \& 5 | * |
| (Vujkovic et al., 2009a) | 1. * <br> 2. Case-control study <br> 3. 161 <br> 4. Netherlands | $\begin{aligned} & \text { 1a. } \\ & 2.195 \\ & 3.22 \end{aligned}$ | 3 | 2a. | * |
| (Vujkovic et al., 2009b) | 1. $*$ <br> 2. Case-control study <br> 3. $*$ <br> 4. Netherlands | $\begin{aligned} & \hline \text { 1a. } \\ & 2.200 \\ & 3.16 \end{aligned}$ | * | 2a. | * |
| (Bastos et al., 2010) | 1. * <br> 2. Case-control study <br> 3. 591 incident Cases of gastric adenocarcinoma and 1463 community controls. <br> 4. Portugal | $\begin{aligned} & 1 \mathrm{a} . \\ & 2.82 \end{aligned}$ | * | 2a. \& 3. \& 4. >1 | * |
| (Bakolis et al., 2010) | 1. Food and Lifestyle and Asthma in Greenwich (F.L.A.G) study <br> 2. Case-control study <br> 3. 1453 individuals ( 599 Cases and 854 controls) <br> 4. UK | 1a. <br> 2.217 <br> 4. Scale of 10: from 'never', to 'two or more times per day') | 1 \& 2. | 1. $>\|0.3\| \& 2 \mathrm{a} . \& 3 . \&$ 5. | * |
| (De Stefani et al., 2008a) <br> (De Stefani et al., 2008b) <br> (De Stefani et al., 2007) | 2. Case-control studies <br> 3. 861 (2008a) <br> 255 Cases and 501 hospitalized controls (2008b) <br> Cases 290 Controls and 290 <br> Cases <br> (2007) <br> 4. Uruguay | 1a. <br> 2.64 <br> 3.27 <br> 4. Scale of 6: never to more than one time per day. | * | $\begin{aligned} & \text { 1. }>0.30 \& 2 \mathrm{~b} . \& 3 . \& \\ & \text { 4. }>1.0 \end{aligned}$ | 7.\& 11 |
| $\begin{aligned} & \text { (De Stefani et al., } \\ & \text { 2004) } \end{aligned}$ | 1. All patients with newly diagnosed and microscopically confirmed gastric carcinomas, admitted for diagnosis and treatment in the four major hospitals <br> 2. Case-control study <br> 3. 240 Cases and 960 controls. <br> 4. Uruguay | $\begin{aligned} & \text { 1a. } \\ & 2.191 \end{aligned}$ | *2 | 2a. \& 3. \& 4 \& 5 | 11. |

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| (Ronco et al., 2006) | 1. <br> 2. $3 .$ <br> 4. | Case -control study 442 newly diagnosed and microscopically confirmed Cases with breast cancer and 442 hospitalized controls Uruguay | $\begin{aligned} & \hline 1 \mathrm{a} . \\ & 2.64 \end{aligned}$ | * | $\begin{aligned} & \text { 1. }>\|0.29\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Wu et al., 2009) | 1. 2. 3. 4. | Case-control study 2172 <br> USA | $\begin{aligned} & \hline 1 \mathrm{a} \\ & 2.174 \end{aligned}$ | * | $\begin{aligned} & \text { 1. }>\|0.30\| \& 2 \mathrm{a} \& 3 . \& \\ & 4 .>1 \& 5 \end{aligned}$ | * |
| (Kubo et al., 2008) | 1. 2. $3 .$ <br> 4. | Case-control study 296 Cases were matched to persons with gastro esophageal reflux disease ( 308 without Barrett's esophagus and to population controls (309 USA | $\begin{aligned} & \text { 1a. } \\ & \text { 2. } 110 \end{aligned}$ | * | $\text { 1. }>0.35 \& 3 . \& 4>1$ <br> \& 5 | * |
| (Murtaugh et al., 2008) | 2. 3. 4. | Four-Corners Breast Cancer (FCBC) study Case-control study (757 Cases, 867 controls) and non-Hispanic white women (1524 Cases, 1598 controls) USA | $\begin{aligned} & \hline \text { 1a. } \\ & 3.69 \end{aligned}$ | * | 1. $>0.35 \& 2 \mathrm{a}$. | * |
| (Dalvi, Canchola \& Horn-Ross, 2007) | 1. 2. 3. 4. | * <br> Case-control study 647 Cases and 633 controls USA | $\begin{aligned} & \hline \text { 1a. } \\ & 2.103 \end{aligned}$ | * | $\begin{aligned} & 1 .>\|0.35\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ | * |
|  <br> Chapman- <br> Novakofski, 2007) | $\begin{aligned} & 2 . \\ & 3 . \\ & 4 . \end{aligned}$ | Health and Nutrition of the Elderly in Botswana. Cross-sectional study 1086 <br> Africa | 1a. <br> 2.21 <br> 4. Scale of 4:"eat less," "more," "same," and "never ate." | * | $\begin{aligned} & \text { 1. }\|>0.5\| 2 a . \& 3 . \& 4 . \\ & >1 . \& 5 . \end{aligned}$ | * |
| (Romaguera et al., 2008) | 1. 2. 3. 4. | Cross-sectional study 1236 Argentina | $\begin{aligned} & \hline 1 \mathrm{a} . \\ & 2.46 \\ & 3.13 \end{aligned}$ | 3 | 1. $>0.2$. \& 4. $>1$ | * |
| (McNaughton et al., 2008) | 2. 3. 4. | Australian National Nutrition Survey (NNS) Cross-sectional study 764 <br> Australia | 1a. <br> 2.127 <br> 4.Scale of 9:"'never or less than once a month" to " 6 or more times per day" |  | $\begin{aligned} & \text { 1. }>0.3 \& 2 . \& 3 . \& 4 \\ & >1.25 \& 5 \end{aligned}$ | 2. \& 6. |

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| (Chen et al., 2006) | 1. | Health Effects of Arsenic <br> Longitudinal Study (HEALS), <br> Cross-sectional study <br> 11116 <br> Bangladesh | $\begin{aligned} & \hline 1 \mathrm{a} . \\ & 2.39 \end{aligned}$ | 1. | $\begin{aligned} & 1 .>\|0.15\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1.5 \& 5 \end{aligned}$ | * |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Custodio das Dores et al., 2007) | 1. 2. 3. 4. | Cross-sectional study 115 Brazil | $\begin{aligned} & \hline \text { 1a. } \\ & 2.97 \end{aligned}$ | 1. | $\begin{aligned} & \text { 1. }>\|0.40\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ | * |
| (Paradis et al., 2009) | 1. 2. 3. 4. | Cross-sectional study. 664 Canada | $\begin{aligned} & 1 . \\ & 2.91 \end{aligned}$ | * | $\begin{aligned} & \text { 1. }>\|0.30\| 2 \mathrm{a} . \& 3 . \& \\ & 4 .>1 \end{aligned}$ | * |
|  <br> Vohl, 2006) | 4. | Cross - sectional study 197 women and 129 men Canada | $\begin{aligned} & 1 \mathrm{a} . \\ & 2.91 \end{aligned}$ | 1. | $\begin{aligned} & 1 .>\|0.30\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ | * |
| (He et al., 2009) | 1. | China National Nutrition and Health Survey Cross-sectional study 56442 <br> China | 1a. | * | 2b. | * |
| (Togo et al., 2004) | 1. | MONICA study (Monitoring of Trends and Determinants in Cardiovascular Diseases). Cross-sectional study 2436 aged $30-60$ y attended all three examinations Denmark | 1a. <br> 2. <br> 3.21 | * | $\begin{aligned} & 1 .>\|0.30\| \& 2 a \& 3 . \& \\ & 4>1 \& 5 \end{aligned}$ |  |
| (Naska et al., 2006) | 4. | Data Food Networking (DAFNE) project Standardized household budget surveys (HBS) Cross-sectional study 94564 households of the ten countries under study, 15251 households whose composition did not fit in any of the predefined Europe | 1e) | * | $\begin{aligned} & \text { 1. }>\|0.20\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ | 3. |
| (Montonen et al., 2005) |  | The Finnish Mobile Clinic Health Examination Survey Cross-sectional study 4304 <br> Finland | $\begin{aligned} & \hline \text { 1a. } \\ & 1.23 \end{aligned}$ | * | $\begin{aligned} & \text { 1. }>\|0.30\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ | 1. \& 7. |
| (Perrin et al., 2005) |  | MONICA study (Monitoring of Trends and Determinants in | $\begin{aligned} & \hline \text { 1c. } \\ & 2.15 \end{aligned}$ | * | $\begin{aligned} & \text { 1. }>\|0.25\| \& 2 a \& 3 . \& \\ & 4>1 \& 5 \end{aligned}$ | 6. \& 7. \& 10. |

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|  |  Cardiovascular Diseases). <br> 2. Cross-sectional study <br> 3. 3508 <br> 4. France |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Kontogianni et al., 2009) | 1. * <br> 2. Cross-sectional study <br> 3. 220 <br> 4. Greece | 1 f . | * | 1 a. $>\|0.3\| \& 2 \mathrm{a}$. | * |
| (Yannakoulia et al., 2008) <br> (Panagiotakos et al., 2007b) | 1. Attica study <br> 2. Cross-sectional study <br> 3. 453 men and 400 women (2008) <br> 1514 men and 1528 women <br> (2007) <br> 4. Greece | $\begin{aligned} & \hline 1 \mathrm{~d} . \\ & 2.156 \\ & 3.22(2007 \mathrm{~b}) \end{aligned}$ | * | 1. $\begin{aligned} & >0.32008 \\ & >\|0.40\| 2007 \\ & \text { 2a } 3 . \& 4>1 \& 5 \end{aligned}$ | 9. \& 10. |
| (Panagiotakos et al., 2007a) | 1. MEDIS (Mediterranean Islands) <br> 2. Cross-sectional study <br> 3. 300 men and women from Cyprus, 142 from Mitilini, 100 from Samothraki, and 104 from Kefalonia islands. <br> 4. Cyprus and Greece | $\begin{aligned} & 1 \mathrm{a} . \\ & 2.15 \end{aligned}$ | * | $\begin{aligned} & \text { 1. }>\|0.3\| \& 2 \mathrm{a} . \& 3 . \& 4 . \\ & >1 . \end{aligned}$ | * |
| (Rezazadeh, Rashidkhani \& Omidvar, 2010) | 1. * <br> 2. Cross-sectional study <br> 3. 460 <br> 4. Iran | $\begin{aligned} & \text { 1a. } \\ & 2.168 \\ & 3.37 \end{aligned}$ | * | 2a. 3. \& 4. >1.0 | * |
| (Esmaillzadeh \& Azadbakht, 2008a) (Esmaillzadeh \& Azadbakht, 2008b) (Esmaillzadeh et al., 2007) | $\begin{array}{ll}\text { 1. } & * \\ \text { 2. } & \text { Cross-sectional study }\end{array}$ <br> 3. 486 <br> 4. Iran | $\begin{aligned} & \text { 1a. } \\ & 2.168 \\ & 3.41 \end{aligned}$ | * | 2a. \& 3. \& $4>1 . \& 5$ | * |
| (Okubo et al., 2008) <br> (Okubo et al., 2007) | 1. $\quad$ * $\quad$ Cross-sectional study <br> 3. 3760 (2008) <br> 3770(2007) <br> 4. Japan | $\begin{aligned} & \hline 1 \mathrm{a} \\ & 2.148 \\ & 3.30 \end{aligned}$ |  | $\begin{aligned} & \text { 1. }>\|0.20\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ | 1. |
| (Sadakane et al., 2008) | 1. * <br> 2. Cross-sectional study <br> 3. 6886 (in the analysis on blood pressure) and 7641 (in the analysis on serum lipids) <br> 4. Japan | 1a. <br> 2.30 <br> 4. Scale of 5: 1: seldom to 5 : almost every day. | * | 2a. \& 3. \& $4>1$. | 1. |
| (Takaoka \& Norback, 2008) | 1. * <br> 2. Cross-sectional study <br> 3. 153 | $\begin{aligned} & \text { 1a. } \\ & 2.11 \end{aligned}$ | * | 2 a . | * |

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|  | 4. | Japan | 4.Scale of five, $0:$ never to 5 : almost daily |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Mizoue et al., 2006) <br> (Mizoue et al., 2005) |  | Self-Defense Forces Health Study Cross-sectional study 2141 Japan | 1a. <br> 2.74 <br> 4.Scale of 7:from <br> "never/, 1 time/mo" to <br> " $2-3$ times/ d." | * | $\begin{aligned} & 1 .>\|0.30\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ | * |
| (Kim et al., 2007) | 1. | Korean National Health and Nutrition Survey Cross-sectional study 1257 <br> Korea | 1 b . | * | * | * |
| (Shin, Oh \& Park, 2007) | 1. | Cross-sectional study 1441 preschool children Korea | 1a. <br> 2.100 <br> 3.31 <br> 4. Scale of 9 :from "rarely use' to 'have three or more times per day" | 1. | 2a \& 3. \& 4>2 \& 5 | * |
| (Craig et al., 2010) | 2. 3. 4. | Cross -sectional study 2352 <br> Scotland | $\begin{aligned} & \hline \text { 1a. } \\ & 2.65 \end{aligned}$ | * | 1. $>\|0.3\| \& 2 \mathrm{a} . \& 3$. | * |
| (Crozier et al., 2008) | 1. 2. 3. 4. | Cross-sectional study 617 women in early pregnancy UK | $\begin{aligned} & \hline \text { 1a. } \\ & 2.100 \\ & 3.49 \end{aligned}$ | * | * | * |
| (Teucher et al., 2007) | 1. | Cross-sectional study $3262$ <br> UK | $\begin{aligned} & \hline \text { 1a. } \\ & 2.131 \\ & 3.54 \end{aligned}$ | 1. \& 2. \& 3 | 3. \& 5 | * |
| (Burt et al., 2006) | 1. 2. 3. 4. | Cross-sectional study $1021$ <br> USA | $\begin{aligned} & \text { 1a. } \\ & 2.20 \end{aligned}$ | * | 3 \& 5 | * |
| (Yang, Kerver \& Song, 2005) | 2. | Cross-sectional study 263 men, 234 women USA | $\begin{aligned} & 1 \mathrm{a} . \\ & 2.22 \end{aligned}$ | 1. | 1. $>0.30$ or $<-0.20 \& 2 \mathrm{a}$ \& 3 . \& $4>1.25 \& 5$ | * |
| (Hughes et al., 2009) | 1. 2. 3 4. | Cohort study. 1119 Australia | 1a. <br> 2.129 <br> 4. Scale of 9:'never"' to " $>1$ times per day.". | * | $\begin{aligned} & \text { 1. }>\|0.15\| 2 \mathrm{a} . \& 3 . \& \\ & 4 .>1 \end{aligned}$ | * |

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| (Muller et al., 2009) | 1. | The Melbourne Collaborative Cohort Study (MCCS) <br> Cohort study 1018 incident prostate cancer Cases <br> Australia | $\begin{aligned} & \hline 1 \mathrm{a} . \\ & 2.121 \end{aligned}$ | * | 4. $>2.0$ | * |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Oddy et al., 2009) |  | The Western Australian Pregnancy Cohort Study Cohort study 1860 Australia | $\begin{aligned} & \hline 1 \mathrm{a} . \\ & 2.1 \\ & 3.26 \end{aligned}$ | * | 1. $>\|0.30\| \& 4 .>1$ |  |  |
| (Shi et al., 2008) |  | Cohort study <br> The total sample included 1308 men and 1541 women, f them 711 participants were from the urban area. <br> China | $\begin{aligned} & 1 \mathrm{a} . \\ & 2.33 \\ & 3.25 \end{aligned}$ | * | $\begin{aligned} & \text { 1.>0.2. \& 2a. \& 3. \& } \\ & 4 . \& 5 . \end{aligned}$ |  |  |
| (Cai et al., 2007) |  | The Shanghai Women's Health Study Cohort study $74942$ <br> China | 1a. <br> 2.71 <br> 4.Scale of 5:daily, weekly, monthly, yearly, or never ) | * | $\begin{aligned} & 1 .>\|0.30\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ |  |  |
| $\begin{aligned} & \text { (Knudsen et al., } \\ & 2008 \text { ) } \end{aligned}$ |  | Danish National Birth Cohort (DNBC) <br> Cohort study 44612 women Denmark | $\begin{aligned} & \hline 1 \mathrm{a} \\ & 3.36 \end{aligned}$ | * | 2a \& 3. \& 4. \& 5. |  |  |
| (Bamia et al., 2007) <br> (Masala et al., 2007) <br> (Pala et al., 2006) <br> (Waijers et al., 2006) | 1. | EPIC-Elderly project Cohort study $\begin{array}{r} 74607(2007 a) \\ 5611(2007 b) \\ 47749(2006) \\ 4990(2005) \end{array}$ <br> Europe | 1a <br> 1d.(2007b) <br> 2.120(2007b) <br> $\quad 188,217$ and <br> $\quad 140(2006 a)$ <br> $\quad 178(2006 b)$ <br> $\quad 148(2005)$ <br> 3. $22(2007 a)$ <br> $\quad 17(2006 b)$ <br> $\quad 57(2006)$ | * | $\begin{aligned} & 1 .>\|0.40\| 2007 a>\|0.30\| \\ & 2007 b \\ & >\|0.20\| 2006 b \& 2 a 3 . \\ & \& 4>1 \& 5 \end{aligned}$ |  |  |
| (Mannisto et al., 2005) <br> (Dixon et al., 2004) | 1. | DIETSCAN project. Three of these Cohort studies (NLCS, ORDET and SMC) who had female participants were included in this specific breast cancer study. | 1a. <br> 2. <br> 51 (2005) <br> 64 (2004) | * | $\begin{aligned} & 1 .>\|0.35\| \& 2 a \& 3 . \& \\ & 4>1 \& 5 \end{aligned}$ |  |  |

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|  | 2. 3. 4. | Cohort study 3123 (2005) 61463 (2004). Europe |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Uusitalo et al., 2009) <br> (Arkkola et al., 2008) | 1. | Finnish Type 1 Diabetes Prediction and Prevention (DIPP) Nutrition Study Cohort study. 3360(2009 study) 3730 (2008 study) Finland | $\begin{aligned} & 1 \mathrm{a} \\ & 2.181 \\ & 3.52 \end{aligned}$ | * | 1. >0.2. \& 2a. \& 3. | * |
| (Mikkila et al., 2005) | 1. 2. 3. 4. | Cohort study. 1200 and 1037 Finland | 1 b . | 1. | $2 \mathrm{a} \& 3 . \& 4>1$ \& 5 | * |
| $\begin{aligned} & \hline \text { (Keskitalo et al., } \\ & 2008 \text { ) } \end{aligned}$ | 1. 2. 3. 4. | Cohort study 2009 <br> Finland | 1a. <br> 2.24 <br> 4. ( 5 categories ( $1=$ never- <br> $5=$ several times a day) | * | 3. \& 4. \& 5. | 2. |
| (Cottet et al., 2009) <br> (Touvier et al., 2009) <br> (Varraso et al., 2009) <br> (Kesse, Clavel- <br>  <br> Boutron-Ruault, 2006) | 1. <br> 2. <br> 3. <br> 4. | E3N [Etude Epidémiologique aupre`s de Femmes dela Mutuelle Générale de l'Education Nationale] study Cohort study <br> 62372(Cottet 2009 study ) <br> 64252(Touvier 2009 study ) <br> 56,881 (Varraso 2009 study ) <br> 516 adenoma Cases ( 175 highrisk adenomas) 4,804 polyp-free women <br> France | 1d. <br> 2.208 <br> 3.57 (Cottet 2009 study ) <br> 46 (Touvier 2009 <br> study ) <br> 56 ( Varraso 2009 <br> study ) <br> 40 ( Kesse 2006 <br> study) | * | 1. <br> $>\|0.25\|$ Cottet 2009 <br> study <br> >0.2 Touvier 2009 <br> study <br> >0.4 Varraso 2009 <br> study <br> 0.20 Kesse 2006 <br> study <br> 2a. <br> 3. <br> 4. $>1.25$ <br> 5. | * |
| (Kim et al., 2008) | 1. 2. 3. 4. | Kohala Health Research Project Cohort Study 1257 participants Hawaii | 1a. <br> 2.166 <br> 4. Scale of 6: ("never," to " 2 times a day or more"). | * | 2a. \& 3. \& 4. >3 | * |
| (Sant et al. 2007) | 1. 2. 3. 4. | ORDET (Hormones and Diet in Etiology of Tumors) Cohort study $8,861$ <br> Italy | $\begin{aligned} & \hline \text { 1a. } \\ & 2.107 \\ & 3.34 \end{aligned}$ | * | $\begin{aligned} & 1 .>\|0.25\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ | 1. \& 7. |
| (Sieri et al., 2004) | 1. | Cohort study | $\begin{aligned} & \hline \text { 1a. } \\ & 2.100 \end{aligned}$ | * | $\begin{aligned} & 1 .>\|0.25\| \& 2 a \& 3 . \& \\ & 4>1 \& 5 \end{aligned}$ |  |

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|  | 3. | $\begin{aligned} & 8984 \\ & \text { Italy } \end{aligned}$ | 3.49 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Nanri et al., 2008) | 1. 2. 3. 4. | Cohort study 3243 men and 4,667 women Japan | 1a <br> 2.49 <br> 4. Scale of 7:(1-3 <br> times/mo, to 3 times/d) | * | 2a. \& 3. \& 4. >1.68 \& $5$ | * |
| (Shimazu et al., 2007) | 1. 2. 3. 4. | ```Ohsaki National Health Insurance (NHI) Cohort study 4 0 5 4 7 Japan``` | 1a. <br> 2.40 <br> 4. Scale of 5:(almost never to almost every day | * | $\begin{aligned} & 1 .>\|0.25\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ | * |
| (Kim et al., 2004) | 1. 2. 3. 4. | Japan Public Health Center (JPHC) study Cohort study 54,498 residents 27,063 men and 27,435. <br> Japan | $\begin{aligned} & 1 \mathrm{a} \\ & 2.60 \end{aligned}$ | 1. | 2a. \& 3. \& 4>1.5 \& 5 | * |
| (Uusitalo et al., 2005) | 1. 2. 3. 4. | NCD Prevention Programme <br> Cohort study <br> 561 men and 554 women <br> Mauritius | 1a. <br> 2.67 <br> 4 Scale of 8: from "never/seldom" to "4 times/d." | * | $2 \mathrm{a} \& 3 . \& 4>1$ \& 5 | * |
| (Balder, Goldbohm \& van den Brandt, 2005) | 1. 2. 3. 4. | The Netherlands Cohort Study on Diet and Cancer Cohort study 58279 <br> Netherlands | $\begin{aligned} & 1 \mathrm{a} \\ & 1 \mathrm{c} . \\ & 2.150 \\ & 3.51 \end{aligned}$ | * | 2a. \& 3. \& $4>1$ \& 5 | * |
| (Brantsaeter et al., 2009) | 1. | Norwegian Mother and Child Cohort Study (MoBa) <br> Cohort study <br> 23423 <br> Norway | $\begin{aligned} & \text { 1a. } \\ & 2.255 \\ & 3.58 \end{aligned}$ | * | $\begin{aligned} & \text { 1. }>\|0.3\| \& 2 \mathrm{a} \& 3 . \& \\ & 4 .>1.6 \end{aligned}$ | * |
| (Engeset et al., 2009) <br> (Engeset et al., 2005) | 1. 2. 3. 4. | NEPIC part of Norwegian <br> Women and Cancer (NOWAC) <br> Cohort study <br> 37 212(2009) <br> 35553 (2005) <br> Norway | $\begin{aligned} & 1 \mathrm{a} . \\ & 2.86 \\ & 3.50 \end{aligned}$ | * |  <br> 4. $>1.5$ | * |
| (Noel et al., 2009) | 1. | Boston Puerto Rican Health Study <br> Cohort study $1167$ <br> Puerto Rico | $\begin{aligned} & \hline \text { 1a. } \\ & 2.126 \\ & 3.34 \end{aligned}$ | 1. | 2a. \& 3. \& 4. | 1. |

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\begin{tabular}{|c|c|c|c|c|c|c|}
\hline (Butler et al., 2008) \& 1. \& \begin{tabular}{l}
Singapore Chinese Health Study Cohort study
\[
63257
\] \\
Singapore
\end{tabular} \& \[
\begin{aligned}
\& \hline \text { 1a. } \\
\& 2.165
\end{aligned}
\] \& * \& \[
\text { 1. }>|0.3| \& 2 \mathrm{a} . \& 3 . \&
\]
\[
5
\] \& * \\
\hline (Butler et al., 2006) \& 1. \& \begin{tabular}{l}
Singapore Chinese Health Study Cohort study \\
623 \\
Singapore
\end{tabular} \& \[
\begin{aligned}
\& \hline \text { 1a. } \\
\& 2.165
\end{aligned}
\] \& * \& \[
\begin{aligned}
\& 1 .>|0.30| \& 2 a \& 3 . \& \\
\& 4>1 \& 5
\end{aligned}
\] \& * \\
\hline (Lopez et al. 2009) \& 4. \& \begin{tabular}{l}
Seguimiento Universidad de \\
Navarra (SUN) study \\
Cohort study \\
11195 \\
Spain
\end{tabular} \& \[
\begin{aligned}
\& \text { 1a. } \\
\& 2.136
\end{aligned}
\] \& * \& \[
\begin{aligned}
\& \text { 1. }>|0.30| \& 2 \mathrm{a} \& 3 . \& \\
\& \text { 4. }>1 \& 5 .
\end{aligned}
\] \& * \\
\hline (Cuco et al., 2006) \& 1. \& Cohort study 11000 Spain \& 1c. \& 3 \& \[
\begin{aligned}
\& \text { 1. }>|0.20| \& 2 \mathrm{a} 3 . \& 4 \\
\& >1 \& 5
\end{aligned}
\] \& * \\
\hline \begin{tabular}{l}
(Akesson et al., 2007) \\
(Newby et al., 2006ba) (Newby et al., 2006b) \\
(Weismayer, Anderson \& Wolk, 2006) \\
(Rashidkhani et al., 2005) \\
(Khani et al., 2004)
\end{tabular} \& 1.
2.
3.

4. \& \begin{tabular}{l}
Swedish Mammography Cohort <br>
Cohort study <br>
24444 (2007) <br>
33840(2006a) <br>
33840(2006b) <br>
66651(2006c) <br>
61431(2005) <br>
66651(2004) <br>
Sweden

 \& 

1a. <br>
2.67 in 1987 and 97 in 1997 <br>
3.26 food groups(2005) <br>
4. Scale of 8: from "never/seldom" to "4 times/d."

\end{tabular} \& \[

$$
\begin{aligned}
& 1 . \\
& 3
\end{aligned}
$$

\] \& \[

$$
\begin{aligned}
& 1>|0.15| \&>|0.20| \& \\
& \text { 2a } 3 . \& 4>1 \& 5
\end{aligned}
$$
\] \& 1. \& 2. \& 7. \& 11. <br>

\hline (Robinson et al., 2009) \& 1.
2.
3.

4. \& | Hertfordshire Cohort Study Cohort study $3217$ |
| :--- |
| UK | \& \[

$$
\begin{aligned}
& 1 \mathrm{~d} . \\
& 3.51
\end{aligned}
$$
\] \& * \& 1.>0.15 \& * <br>

\hline (Akbaraly et al., 2009) \& 1. \& | Whitehall II study Cohort study 3486 |
| :--- |
| UK | \& | 1a. |
| :--- |
| 2.127 |
| 3.37 |
| 4. Scale of 9: from 'never, or less than once per month' to 'six or more times per day'.) | \& 1. \& * \& * <br>


\hline (Ambrosini et al., 2009) \& 4. \& | Raine Study |
| :--- |
| Cohort study 2900 |
| Australia | \& \[

$$
\begin{aligned}
& \hline \text { 1a. } \\
& 2.212 \\
& 3.38
\end{aligned}
$$

\] \& * \& \[

$$
\begin{aligned}
& \text { 1. }>|0.30| \& 2 \mathrm{a} \& 3 . \& \\
& 4 .>1 . \& 5
\end{aligned}
$$
\] \& * <br>

\hline
\end{tabular}

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| (Hamer \& Mishra, 2010) |  | Low Income Diet and Nutrition Survey (LIDNS) <br> Cohort study <br> 3728 <br> UK | $\begin{aligned} & \hline 1 \mathrm{~b} . \\ & 2.51 \end{aligned}$ | * | 1. $>\|0.30\| \& 2 \mathrm{a}$. | 2. |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Shaheen et al., 2009) <br> (Wiles et al., 2009) <br> (Feinstein et al., 2008) <br> (Northstone, <br> Emmett \& Rogers, 2008a) <br> (Northstone, <br> Emmett \& Rogers, 2008b) <br>  <br> Emmett, 2008b) <br>  <br> Emmett, 2005) |  | Avon Longitudinal Study of Parents and Children (ALSPAC) Cohort study <br> 14062(shaheen 2009) <br> 951(Wiles 2009) <br> 13988(Feinstein 2008) <br> 14541 (Northstone 2008a ) <br> 6271 (Northstone 2008b ) <br> 12035 ((Northstone and Emmet 2008b) <br> 8515 ((Northstone \& Emmett, 2005) <br> UK | 1a. <br> 2. 90 <br> 3. 43 (Shaheen 2009) <br> 34 to 41 (Wiles 2009) <br> 43 to 54 (Feinstein 2008) <br> 44 (Northstone 2008a ) <br> 44(Northstone 2008b) <br> 44(Northstone and <br> Emmet 2008b) <br> (Northstone 2005) <br> 4. Scale of 5: from never or rarely to more than once a day | * | $\begin{aligned} & \text { 1. }>\|0.30\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ | * |  |
| (Borland et al., 2008) <br> (Robinson et al., 2007) <br> (Crozier et al., 2006) <br> (Robinson et al., 2004) |  | Southampton Women's Survey (SWS) study Cohort study 6129 (2008 study) 434 (2007 study) 617 (2006 study) 3779 (2004 study) UK | $\begin{aligned} & \text { 1a. } \\ & 2.100 \\ & 3.49 \end{aligned}$ | 3 | $\begin{aligned} & 1 .>\|0.20\| \& 2 a \& 3 . \& \\ & 4>1 \& 5 \end{aligned}$ | * |  |
| (McNaughton et al., 2007) |  | The Medical Research Council (MRC) National Survey of Health and Development (NSHD, also known as the 1946 British Birth Cohort) <br> Cohort study <br> 1265 <br> UK | $\begin{aligned} & \hline 1 \mathrm{c} . \\ & 3.126 \end{aligned}$ |  | $\begin{aligned} & 1 .>\|0.25\| \& 2 \mathrm{a} \& 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ | * |  |
| $\begin{aligned} & \text { (Agurs-Collins et al., } \\ & 2009 \text { ) } \end{aligned}$ |  | Black Women's Health Study (BWHS) <br> Cohort study 1144 incident Cases USA | 1a. <br> 2.69 <br> 3.29 <br> 4.Scale of 9 : from "never or 1 per month" to " 2 or more per day" for each food |  | $\begin{aligned} & \text { 1. }>\|0.40\| \& 2 \mathrm{a} . \& 3 . \& \\ & 4 \& 5 . \end{aligned}$ | * |  |

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| (Cutler et al., 2009) |  | Project EAT study <br> Cohort study <br> At time 1, 4746. At Time 2, 2516 participants USA | $\begin{aligned} & \hline 1 \mathrm{a} . \\ & 2.152 \end{aligned}$ | * | 2a. \& 3. \& 4.>1 | * |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| (Deshmukh-Taskar et al., 2009) |  | Bogalusa Heart Study (BHS) Cohort study 995 USA | 1a. <br> 2.31 <br> 4. Scale of 6: from 'never or less than once a month' to 'five or more times per day.' | 1. | $\begin{aligned} & \text { 1. }>\|0.30\| \& 3 . \& 4 . \\ & >1 . \end{aligned}$ | * |  |
| (Erber et al., 2010) (Takata et al., 2007) (Park et al., 2005) | 2. | ```Multiethnic Cohort (MEC) study Cohort study 36256 men and 39256 women (2010 study ) 3512 (2007 study ) 195278 (2005 study) USA``` | $\begin{aligned} & 1 \mathrm{a} . \\ & 2.200 \end{aligned}$ | * | $\begin{aligned} & 1 .>\|0.60\| \& 2 a . \& 3 . \& \\ & 4>1.25 \& 5 \end{aligned}$ | 7 |  |
| $\begin{aligned} & \text { (Imamura et al., } \\ & \text { 2009) } \end{aligned}$ |  | Framingham Offspring Study Cohort study $2,879$ <br> USA | 1a. <br> 2.126 <br> 3.40 <br> 4.Scale of 9 | * | 3. \& 4. > 1.0 | * |  |
| (Qi et al., 2009) <br> (Varraso et al., 2007b) <br> (Wu et al., 2006) <br> (Michaud et al., 2005) <br> (Wu et al., 2004a) | 1. <br> 2. <br> 3. <br>  <br> 4. | The Health Professionals Follow-Up Study (HPFS) Cohort study <br> 1996 Cases with 1337 controls (2009 study) <br> 42917 (2007 study) <br> 47725 (2006 study) <br> 51529(2005 study) <br> 20888 (2004 study) <br> USA | 1a. <br> 2.131 <br> 3.40 <br> 4. Scale of 9: from "almost never" to "6 times/d." | * | $\begin{aligned} & \text { 1. }>\|0.30\| \& 2 \mathrm{a} 3 . \& 4 \\ & >1 \& 5 \end{aligned}$ | * |  |
| (Lutsey et al., 2009) |  | Iowa's Women Health Study (IWHS) <br> Cohort study $1950$ <br> USA | 1a. <br> 2.127 <br> 3.39 <br> 4. Scale of 9: from never or 1 serving per month to $\geq 6$ servings per day. | * | 2a. \& 4. $>2.0$ \& 5. | * |  |
| (Nettleton et al., 2009) <br> (Nettleton et al., | 2. | Multi-Ethnic Study of Atherosclerosis( MESA ) Cohort study | $\begin{aligned} & \text { 1a. } \\ & 2.120 \\ & 3.47 \end{aligned}$ | 1. | 2a. \& 3. \& $4>1$ \& 5 | 8. |  |

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\begin{tabular}{|c|c|c|c|c|c|c|}
\hline \begin{tabular}{l}
2008b) \\
(Nettleton et al., 2008a) \\
(Nettleton et al., 2007)
\end{tabular} \& 3.
4. \& \[
\begin{aligned}
\& 5316 \text { (2009 study ) } \\
\& 5011 \text { (2008b study) } \\
\& 5042 \text { (200a study) } \\
\& 5089 \text { (2007 study) } \\
\& \text { USA }
\end{aligned}
\] \& 4. Scale of 9: from "rare or never" to a maximum of " 2 times per day \& \& \& \\
\hline \begin{tabular}{l}
(Reedy et al., 2010) \\
(Flood et al., 2008)
\end{tabular} \& \begin{tabular}{l}
1. \\
2. \\
4.
\end{tabular} \& \begin{tabular}{l}
National Institute of Health NIH-AARP Diet and Health Study, Cohort study \\
492306 (2009 study) \\
49282 (2008 study) \\
USA
\end{tabular} \& \begin{tabular}{l}
1a. \\
2. 204 \\
4. Scale of 10 : "never" to " 6 times/d" for beverages and from "never" to " 2 times/d" for solid foods as
\end{tabular} \& 1. \& 2. \& 3. \& 2a. \& 3. \& 5. \& * \\
\hline (Chang et al., 2008) \& 2.
3
4. \& \[
\begin{aligned}
\& \text { California Teachers Study } \\
\& \text { Cohort study } \\
\& 311 \\
\& \text { USA } \\
\& \hline
\end{aligned}
\] \& \[
\begin{aligned}
\& \hline 1 \mathrm{a} . \\
\& 2.112
\end{aligned}
\] \& * \& 2a. \& 3. \& 4. \& 5. \& * \\
\hline \begin{tabular}{l}
(Heidemann et al., 2008) \\
(Varraso et al., 2007a) \\
(Schulze et al., 2006) \\
(Zhang et al., 2006) \\
(Fung et al., 2005) \\
(Kroenke et al., 2005) \\
(Fung et al., 2004a) \\
(Fung et al., 2004b) \\
(Lopez-Garcia et al., 2004)
\end{tabular} \& 1.
2.
3.

4. \& NHS I and II (Nurses' Health
Study )
Cohort study
$\mathbf{7 2 1 1 3 ( 2 0 0 8 )}$
$51670(2006 a)$
$13110(2006 b)$
$71058(2005 a)$
2619(2005b)
69554(2004a)
71768(2004b)
121700(2004c)
72043(2007)

USA \& | 1a. |
| :--- |
| 2.116 and 133 |
| 3.37 to 39 |
| 4. Scale of 9: from "never or less than once per month" to "6 or more times per day. | \& * \& \[

$$
\begin{aligned}
& \text { 2a. \& 3. \& 4. > 1, } \\
& >2.75 \& 5 .
\end{aligned}
$$
\] \& 11 <br>

\hline (Lutsey, Steffen \& Stevens, 2008) \& 1
2
2
3

4 \& Atherosclerosis Risk in Communities (ARIC) study Cohort study 3782 incident Cases USA \& | 1a. |
| :--- |
| 2.66 |
| 3.32 |
| 4 Scale of 9: from never or 1 time a month to 6 times a day. | \& * \& \[

$$
\begin{aligned}
& \text { 1. >0.2. \& 2. \& 3. \& } 4 \\
& >2.0 \& 5
\end{aligned}
$$
\] \& * <br>

\hline (Tseng et al., 2008) \& \& | Minnesota Breast Cancer Family |
| :--- |
| Study |
| Cohort study |
| 3147 |
| USA | \& | 1a. |
| :--- |
| 2. 153 |
| 4. Scale of 8: from "never or less than once per month"' to "six or more times per day." | \& 1. \& \[

$$
\begin{aligned}
& \text { 1. }>0.20 \& 2 . \& 3 . \& 4 . \\
& >1 .
\end{aligned}
$$
\] \& 7. <br>

\hline
\end{tabular}

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| (Meyerhardt et al., 2007) | 1. NCI sponsored Cancer and Leukemia Group B (CALGB) <br> 2. Cohort study <br> 3. 1009 <br> 4. USA | 1a. <br> 2.131 <br> 3.39 <br> 4. Scale of 9: from never to 6 or more times per day. | * | 4.>1.5 \& 5. | * |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Velie et al., 2005) | 1. Breast Cancer Detection <br> Demonstration Project (BCDDP) <br> 2. Cohort study <br> 3. 280000 <br> 4. USA | $\begin{aligned} & \hline 1 \mathrm{a} \\ & 2.61 \end{aligned}$ | 1 \& 2 \& 3 | $\begin{aligned} & \text { 1. }>\|0.20\| \& 2 \mathrm{a} \& 3 . \& \\ & 4>1 \& 5 \end{aligned}$ | * |
| (Newby et al., 2004) <br>  <br> Tucker, 2004 | 1. Baltimore Longitudinal Study (BLS) <br> 2. Cohort study <br> 3. 459 <br> 4. USA | $\begin{aligned} & 1 \mathrm{a} . \\ & 2.100 \\ & 3.49 \end{aligned}$ | * | $\begin{aligned} & 1 .>\|0.25\| \& 2 \mathrm{a} \& 3 \& \\ & 4>1 \& 5 \end{aligned}$ | * |
| (Tseng et al., 2004) | 1. NHANES <br> 2. Cohort study <br> 3. 136 <br> 4. USA | $\begin{aligned} & 1 \mathrm{a} \\ & 2.26 \end{aligned}$ | * | $\begin{aligned} & \text { 1. }>\|0.20\| \& 2 \mathrm{a} \& 3 . \& \\ & 4>1 \& 5 \end{aligned}$ |  |
| (Lau et al., 2008) | 1. Inter99 study <br> 2. population based randomised intervention study <br> 3. Baseline data of 3372 women and 3191 men (30-60 years old) from the population-based survey Inter99 was used. <br> 4. Denmark | $\begin{aligned} & \hline 1 \mathrm{a} . \\ & 2.198 \\ & 3.34 \end{aligned}$ | * | 1. $>0.40$ \& 2 b . | 1. \& 11. |
| (Cottet et al., 2005) | $\begin{array}{ll}\text { 1. } & \begin{array}{l}\text { European Cancer Prevention } \\ (\mathbf{E C P})\end{array} \\ \text { 2. } \begin{array}{l}\text { Intervention Study } \\ \text { a randomized trial of calcium } \\ \text { and fiber supplementation }\end{array} \\ \text { 3. } \\ \text { 277 men and } 165 \text { women } \\ \text { 4. } & \text { Europe }\end{array}$ | $\begin{aligned} & 1 \mathrm{a} . \\ & 3.50 \end{aligned}$ | * | $\begin{aligned} & \text { 1. }>\|0.25\| \& 2 a \& 3 . \& \\ & 4>1 \& 5 \end{aligned}$ | 6. |
| $\begin{aligned} & \text { (Kesse-Guyot et al., } \\ & \text { 2009) } \end{aligned}$ | 1. SU.VI.MAX (Supple'mentation en Vitamines et Mine'rau Antioxydants) study <br> 2. a randomized,double-blind, placebo-controlled primary prevention trial <br> 3. 2463 women and 2731 men. <br> 4. France | 1 b . | * | $\begin{aligned} & \text { 1. }>\|0.30\| \& 2 \mathrm{a} \& 3 . \& \\ & \text { 4. }>1 . \& 5 \end{aligned}$ | 1. |
| (Murtaugh et al., | 1. * <br> 2. Population-based control | $\begin{aligned} & 1 \mathrm{a} . \\ & 3.68 \end{aligned}$ | * | $\begin{aligned} & 1 .>\|0.35\| \& 2 \mathrm{a} \& 3 . \& 4 \\ & >2 \& 5 \end{aligned}$ | 3. \& 8. |

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| 2007) | participants <br> 3. <br> 871 Hispanic and 1599 non- <br> Hispanic |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 4. | USA |  |  |  |  |

*Information are not provided

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Table C. Number of dietary patterns, label of dietary patterns, foods that correlated highly with empirically derived dietary patterns in each study. Furthermore associations of dietary patterns with disease outcomes and socio-demographic characteristics are presented (Sorted alphabetically by health outcome).

| Author | Investigated health outcome in each study | Number of dietary patterns, label of dietary patterns and foods that correlated highly with the dietary patterns <br> 1. Number of dietary patterns (median: 3, IQR: 2-4) <br> 2. Label of each pattern/foods and food groups that correlated highly with this pattern <br> 3. Percentage of total variance of original food items being explained by the dietary patterns in each study. (median: $24 \%$, IQR: 19.9-31.3). | Effect of dietary pattern on investigated health outcomes in each study <br> Dietary pattern label: (Effect estimate , 95\% CI, P for trend) | Direction of association of dietary patterns with socio-demographic variables in each study |
| :---: | :---: | :---: | :---: | :---: |
| $\begin{aligned} & \text { (Erber et al., } \\ & 2010 \text { ) } \end{aligned}$ | Type 2 Diabetes | 1. 3 <br> 2. Fat and Meat: fat, meat, eggs, and cheese Vegetables : vegetables and also fruits Fruit and Milk: milk, yogurt, cheese, and fruits <br> 3. $30 \%$ | Fat and Meat: significantly associated with diabetes risk in men ( $\mathrm{HR}=1.40 ; 95 \% \mathrm{CI}: 1.23-1.60 ; \mathrm{P}$ for trend $<0.0001$ ) and women ( $\mathrm{HR}=1.22 ; 95 \% \mathrm{CI}: 1.06-1.40 ; \mathrm{P}$ for trend=$=0.004$ ) when comparing extreme quintiles) <br> Other patterns : Not Statistically Significant | * |
| (Imamura et <br> al., 2009) | Type 2 Diabetes | 1. 3 <br> 2. Western: High fat dairy, eggs , processed meat, refined grains, meat, potato, fried foods <br> Prudent: reduced fat dairy , fruits, vegetables, sweet baked goods Alcohol : white wine, fruits, vegetables ,fish and vegetables | Alcohol: inversely associated with T2D risk ( $\mathrm{HR}=0.33$, 95\% CI: 0.17-0.64) <br> Other patterns : Not Statistically Significant | * |
| $\begin{aligned} & \text { (Qi et al., } \\ & 2009) \end{aligned}$ | Type 2 Diabetes | 1. 2 <br> 2. Prudent: vegetables, fruit, legumes, whole grains, fish, and poultry. <br> Western: processed meat, red meat, butter, high-fat dairy products, eggs, and refined grains | Western: A significant interaction ( $\mathrm{P}<0.02$ ) was observed between the T2D. The multivariable odds ratios (ORs) of T2D across increasing quartiles for the Western dietary pattern were $1.00,1.23$ ( $95 \%$ CI: 0.88-1.73), 1.49 (1.062.09), and 2.06 (1.48-2.88) among men with a high GRS (12 risk alleles; P for trend $<0.01$ ). <br> Other patterns : Not Statistically Significant | Prudent: positively associated with older age, physical activity, and inversely associated with smoking <br> Western : positively associated with drinking and smoking |
| $\begin{aligned} & \text { (Kim et al., } \\ & 2008) \end{aligned}$ | Type 2 Diabetes | 1. 3 <br> 2. Factor 1: fruits, vegetables, and bean products. <br> Factor 2 : corned beef and cabbage, rice, | Factor 2: positively associated with T2DM (OR 1.30, 95\% CI 1.03-1.68) | Factor 1 : negatively correlated with BMI, smoking, and positively correlated with years of education and physical activity <br> Factor 2 : positively correlated with BMI, smoking, |

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|  |  | steamed shell fish, Filipino ethnic foods and local Hawaiian foods <br> Factor 3: French fries, fast-foods hamburgers, pizza, chips, soda, pasta, and salad dressings. |  | and negatively correlated with income level, years of education, physical activity <br> Factor 3 : positively correlated with income and education and negatively correlated with BMI |
| :---: | :---: | :---: | :---: | :---: |
| (McNaughton et al., 2008) | Type 2 Diabetes | 1. 3 <br> 2. Fruit, salad, cereals, and fish High fat and sugar Vegetables <br> 3. $11.9,5.9$, and $3.9 \%$ | Fruit, salad, cereals, and fish : inversely associated with diastolic blood pressure ( $\mathrm{P}<0.001$ ) <br> High fat and sugar: associated with increased risk of type 2 diabetes (HR top quartile 2.95 ( $95 \% \mathrm{CI}: 2.19-3.97$ ); adjusted for age, sex, and energy misreporting). This relationship was attenuated after adjustment for ethnicity, employment grade, health behaviours (smoking, alcohol use, and physical activity) but remained significant after further adjustment for blood pressure and BMI (HR: 1.51 ( $95 \% \mathrm{CI}$ : 1.10-2.09)). <br> Other patterns: Not Statistically Significant | High fat and sugar : positively associated with being male <br> Vegetables : positively associated with rural region of residence <br> Fruit, salad, cereals, and fish : inversely associated with age |
| $\begin{aligned} & \text { (Nanri et al., } \\ & 2008 \text { ) } \end{aligned}$ | Type 2 Diabetes | 1. 4 <br> 2. Healthy: vegetables, fruit, soy products, fish, and yogurt. <br> High-fat: meat, processed meat, mayonnaise, and egg <br> Seafood: seafood, shellfish, salted fish guts, fish roe, and fish-paste products Westernized breakfast: bread, margarine, and coffee and low intakes of rice and miso soup <br> 3. $30.5 \%$ | Westernized breakfast: inversely related to A1C concentrations ( P for trend 0.02 in both men and women); the multivariate-adjusted ORs for the highest versus lowest quintiles were $0.60(95 \% \mathrm{CI}: 0.43-0.84)$ and $0.64(95 \% \mathrm{CI}$ : $0.46-0.90$ ) for men and women, respectively. <br> Seafood : positively associated with A1C concentrations in men only ( P trend 0.01 ) <br> Other patterns: Not Statistically Significant | Healthy: positively correlated with older age, physical activity, leisure time, and negatively associated with smoking status and alcohol drinking. High-fat: associated positively with younger age and smoking. <br> Seafood: positively associated with alcohol drinking and higher BMI. <br> Westernized breakfast: positively associated with younger age, lower BMI, no smoking status and physical activity |
| (Nettleton et al., 2008a) | Type 2 Diabetes | 1. 5 <br> 2. Fats and processed meats <br> Vegetables and fish beans, Tomatoes Refined grains <br> Whole grains and fruit | Vegetables and fish beans, Tomatoes <br> Refined grains: associated with an 18\% greater risk (HR per 1-score SD 1.18 ( $95 \%$ CI $1.06-1.32$ ); P trend 0.004 ) Whole grains and fruit associated with a $15 \%$ lower diabetes risk (HR: 0.85 [0.76-0.95]; P- trend 0.005 . | * |
| (Montonen et al., 2005) | Type 2 Diabetes | 1. 2 <br> 2. Prudent: fruits and vegetables Conservative: butter, potatoes, and whole milk. | Prudent : associated with reduced risk of type 2 diabetes (RR 0.72 ( $95 \%$ CI: 0.53-0.97; p for trend $=0.03$ ) <br> Conservative : associated with increased risk of type 2 diabetes (RR 1.49 ( $95 \%$ confidence interval: 1.11, 2.00; p for trend $=0.01$ ) | * |
| $\begin{aligned} & \text { (Fung et al., } \\ & \text { 2004a) } \end{aligned}$ | Type 2 Diabetes | 1. 2 <br> 2. Prudent: fruits, vegetables, whole grains, fish, poultry, and low-fat dairy Western: red and processed meats, refined grains, sweets and desserts, and high-fat | Western: associated with increase drisk of type 2 diabetes RR for type 2 diabetes of $1.49(95 \%$ confidence interval $[\mathrm{CI}], 1.26-1.76, P$ for trend, $<.001$ ) when comparing the highest to lowest quintiles of the Western pattern | * |

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|  |  | dairy products. | Other patterns: Not Statistically Significant |  |
| :---: | :---: | :---: | :---: | :---: |
| (Varraso et al., 2007a) | Respiratory and Allergic symptoms (COPD) | 1. 2 <br> 2. Prudent: fruit (fresh apples or pears; oranges; peaches, apricots, or plums; strawberries; cantaloupes; blueberries; grapefruits), vegetables (broccoli, eggplant, cauliflower, coleslaw, carrots, raw spinach, celery, string beans, romaine leaf lettuce, yellow squash, cooked spinach, iceberg head lettuce, tomatoes, mushrooms, Brussels sprouts, mixed vegetables, garlic, beans lentils, beets), poultry (chicken or turkey without skin), and fish. <br> Western: French fries, hamburger, cured meats (processed meats, hot dogs, bacon), sweets and desserts (home-baked cake, doughnuts, brownies, ready-made sweet rolls, home-baked pies, pancakes or waffles), and refined cereals (white bread, pasta) | Prudent : negatively associated with risk of newly diagnosed COPD (relative risk (RR) for highest compared with lowest quintile: $0.75 ; 95 \% \mathrm{CI}: 0.58-0.98$; P for trend 0.02) <br> Western: positively associated with risk of COPD (RR for highest compared with lowest quintile: $1.31 ; 95 \%$ CI: $0.94-$ 1.82; P for trend 0.02 . | Prudent: positively associated with physically activity lower body mass index, non-smoking, and being a female <br> Western: positively associated with being men, white, and current smoker. |
| (Butler et al., 2006) | Respiratory and Allergic Symptoms (new cough in phlegm) | 1. 2 <br> 2. Vegetable-fruit-soy: vegetable, fruit, and soy food intake; of the 32 foods included in the pattern, 23 were vegetables, five were soyfood items, and four were fruit items. Meat-dim sum: chicken, pork, fish, rice and noodle dishes, and preserved foods | Meat-dim sum: positively associated with new-onset cough with phlegm (OR, 1.43; 95\% CI: 1.08-1.89, p for trend 0.02 ). <br> Other patterns: Not Statistically Significant | Vegetable-fruit-soy: was positively correlated with people who were less likely to, to smoke, and to lack formal education. <br> Meat-dim sum : : was positively correlated with people who were more likely to be younger, male, current smokers, and have had formal education |
| (Hooper et <br> al., 2008) | Respiratory and allergic symptoms (asthma, atopy) | 1. 2 <br> 2. Urban/rural dimension: Foods which were strongly negatively correlated with the (characteristic of rural diets) were pumpkin leaves, young pumpkin, cabbage and wild leaves and berries. Foods which were strongly positively associated with the urban/rural dimension of, diet (characteristic of urban diets) were fried potatoes, carrots, tinned fruit salad, chicken, sausages, yoghurt, packet custard and jelly. Second convenience rotated dimension <br> 3. $25 \%$ | The Urban component of diet was strongly associated with positive skin tests even after adjusting for urban residence (OR: 2.1;95CI: 1.2-3.7, P-value= 0.009 ) <br> Other patterns: Not Statistically Significant | * |

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\begin{tabular}{|c|c|c|c|c|c|}
\hline \begin{tabular}{l}
(Bakolis et \\
al., 2010)
\end{tabular} \& Respiratory and Allergic Symptoms (Asthma) \& 1.
2.

3. \& \begin{tabular}{l}
5 <br>
Prudent: wholemeal bread and rolls, yoghurt, cheese, fish, salad vegetables, pasta, couscous, vegetable dishes and different types of dressing. Vegetable and Fruit: vegetables and fruits Western: white bread and rolls, chips, roast potatoes, baked beans, processed meats, bacon, ham, crisps, meat dishes, fried snacks, chocolate bars, sponge puddings and cakes, ketchup, coke. <br>
Traditional: vegetables, pork, beef, liver and lamb, and a low intake of naan paratha and Bombay mix. <br>
Vegetarian: cream crackers, cre`me fraiche, macaroni cheese, chick peas, houmous, lentils, nut roast, vegetables, nuts and seeds $25 \%$

 \& 

Vegetarian: positively associated with asthma [adjusted odds ratio comparing top vs. bottom quintile of pattern score ( $\mathrm{OR}=1.43,95 \% \mathrm{CI}: 0.93-2.20$ ), P trend 0.075 ). <br>
Traditional: associated negatively with asthma [OR $=0.68$, 95\%CI: 0.45-1.03., P trend< 0.071). <br>
Prudent: positively associated with chronic bronchitis (OR= 2.61 ( $95 \% \mathrm{CI}: 1.13-6.05, \mathrm{P}$ trend 0.025 ). <br>
Other patterns : Not Statistically Significant
\end{tabular} \& * <br>

\hline | (Hooper et |
| :--- |
| al., 2010) | \& Respiratory and allergic symptoms (asthma) \& 1.

2. 
3. \& | 2 |
| :--- |
| Meat and potato: meat, beef, pork, bacon, sausage and fried egg/scrambled egg/omelette intake at all the centres and also with intake of potato or chips, bread, butter, biscuits and cakes. |
| Fish, fruits and vegetables: several fruits and less consistently with intakes of a number of vegetables and fish. $11.2 \%$ | \& All the patterns : Not Statistically Significant \& * <br>

\hline | (Shaheen et |
| :--- |
| al., 2009) | \& | Respiratory and Allergic |
| :--- |
| Symptoms (Asthma) | \& 1.

2. 
3. \& \begin{tabular}{l}
5 <br>
Health conscious: salad, fruit, fruit juices, rice, pasta, oat/bran based breakfast cereals, fish, pulses, cheese, non-white bread Traditional: vegetables, red meat, poultry Vegetarian: meat substitutes, pulses, nuts, herbal tea <br>
Processed: meat pies, sausages, burgers, fried foods, pizza, chips, crisps, white bread, eggs, baked beans <br>
Confectionery: chocolate, sweets, biscuits, cakes, puddings
(31.6\%)

 \& * \& 

Health conscious: positively associated with education, age and non-white women and negative associated with parity, being single, non-working women, smokers and overweight before prepregnancy. <br>
Processed: Opposite associations
\end{tabular} <br>

\hline (Varraso et al., 2009) \& Respiratory and Allergic Symptoms (Asthma) \& 1. \& | 3 |
| :--- |
| Prudent: fruits and vegetables |
| Western: pizza/salty pies, dessert, cured meats and pasta | \& Western: associated with an increased risk of reporting frequent asthma attacks (highest versus lowest tertile odds ratio (OR) 1.79, 95\%confidence interval (CI) 1.11-3.73). \& * <br>

\hline
\end{tabular}

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|  |  | Nuts and Wine : nuts and seeds, salty biscuits, olives, wine, and fortified wine | Nuts and Wine: was associated with a decreased risk of reporting frequent asthma attacks (highest versus lowest tertile OR $0.65,95 \%$ CI $0.31-0.96$ ). <br> All the patterns: Not Statistically Significant |  |
| :---: | :---: | :---: | :---: | :---: |
| (Varraso et al., 2007b) | Respiratory and allergic symptoms ( COPD) | 1. 2 <br> 2. Prudent: fruits, vegetables, fish, poultry and whole grain products Western: refined grains, cured and red meats, desserts and sweets, French fries, eggs and high-fat dairy products. | Prudent : inversely associated with the risk of newly diagnosed COPD (RR for highest vs. lowest quintile 0.50 ( $95 \%$ CI $0.25-0.98$, p for trend $=0.02$ ) <br> Western : positively associated with the risk of newly diagnosed COPD (RR for highest vs. lowest quintile 4.56 ( $95 \%$ CI: $1.95-10.69$, p for trend $<0.001$ ) | Prudent :positively associated with people who were more physically active, less likely to be current smokers and took more multivitamin supplements, Western : positively associated with higher BMI, less physical activity, less smoking and fewer multivitamin supplements consumption |
| (Takaoka \& Norback, 2008) | Respiratory and Allergic Symptoms (wheeze and respiratory infections) | 1. 5 <br> 2. The first factor: fruit, raw vegetable and cooked vegetable <br> The second factor : fast food, soft drink and juice <br> The third factor : meat, fish and seafood The fourth factor : milk and yoghurt consumption <br> The fifth factor : butter and rapeseed oil | Second factor (fast food, juice and soft drinks) : positively related to wheeze and respiratory infections (OR:1.19 95\%CI: (1.04-1.37), P trend $<0.01$ ) <br> Fifth Factor : positively related to wheeze and respiratory infections (OR:2.17 (1.31-3.59) P trend $<0.003$ ) |  |
| (Bamia et al., 2007) | Overall Mortality | $\begin{array}{ll}\text { 1. } & 1 \\ \text { 2. plant-based }\end{array}$ <br> 3. $14.6 \%$ | Plant-based: associated with a lower overall mortality, a one standard deviation increment corresponding to a statistically significant reduction of $14 \%$ ( $95 \%$ confidence interval 5-23\%). | * |
| $\begin{aligned} & \text { (Masala et al., } \\ & \text { 2007) } \end{aligned}$ | Overall mortality | 1. 4 <br> 2. Prudent: cooked vegetables, legumes, fish, and seed oil <br> Pasta \& Meat: pasta and other grains, tomato sauce, red and processed meats, added animal fat, white bread and wine; on the other hand, this pattern showed a low consumption of yoghurt. <br> Olive Oil \& Salad: olive oil as added fat, raw vegetables (tomatoes, leafy and root vegetables), soups and white meat (chicken and turkey) <br> Sweet \& Dairy: sugar, cakes, ice-cream, coffee, eggs, butter, milk and cheese. <br> 3. $21 \%$ | Olive Oil \& Salad: inversely associated with overall mortality. After adjustment for gender, age and caloric intake, overall mortality was reduced by approximately $50 \%$ in the highest quartile and a significant trend emerged ( $\mathrm{P}<0 \cdot 008$ ). <br> Other patterns: Not Statistically Significant | Prudent: positively associated with people who were more frequently females, more educated, more likely to be single, former smokers and obese Pasta \& Meat: positively associated with people who were more likely among married males, current smokers, overweight or obese subjects. <br> Olive oil \& Salad: positively associated with people who were more frequently males, married, with a higher school education, leaner and more physically active. <br> Sweet \& Dairy: positively associated with people who were more likely to have a higher education, to be more physically active and with a normal weight. |
| (Waijers et al., 2006) | Overall mortality | 1. 3 <br> 2. Mediterranean-like: pasta and rice, sauces, fish, and vegetables in combination with | Healthy Traditional was associated with a lower mortality rate (Women in the highest tertile of this pattern had ab30\% | Mediterranean-like: positively associated with people who were younger, higher educated, and |

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|  |  | ( | vegetable oils, wine, and other cereals. Foods such as potatoes, bread, and margarine, contributed negatively to this component. <br> Traditional Dutch: meat, potatoes, vegetables, eggs and alcoholic beverages. It was low in intakes of dairy products, sweets, and pastries. <br> Healthy Traditional: vegetables, fruit, dairy products, potatoes, and legumes, and also non-alcoholic beverages. It was low in intakes of butter and alcoholic beverages 25\% | lower mortality risk than those in lowest tertile (95\% CI for the hazard ratio: 0.52-0.95) <br> Other patterns : Not Statistically Significant | more often former smokers <br> Traditional Dutch: positively associated with people who had a lower level of education, were more current smokers, and were more overweight. Healthy Traditional: positively associated with people who were less educated, more likely nonsmokers, had higher BMIs, and were more physically active |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Vujkovic et <br> al., 2009b) | Other(Spina Bifida) | 2. | 1 <br> Mediterranean | Mediterranean: positively associated with increased risk of spina bifida of offspring in mothers (OR=3.5 ( $95 \% \mathrm{CI}$ : 1.5-7.9). | Mediterranean: positively associated with higher maternal age at birth of the index child, higher education and more alcohol consumption in the preconception period |
| (D'Souza et <br> al., 2008) | Other(Chron's Disease) |  | 4 <br> Traditional Western (girls): meat, fried food items, fast foods, snacks, and desserts Prudent (girls): vegetables, fruits, dairy products, eggs, olive oil, dark breads, grains, fish, and nuts <br> Cheese-Snack (girls): cheese, snacks, and desserts and this. <br> Beverage (girls): beverages (tea, coffee, coke, and milk-shakes), some organ meats, and salsa <br> Partial Western (boys): beverages, fast foods, snacks, white bread, meat sandwiches, and dessert items. <br> Prudent(boys): vegetables, fruits, yogurt, olive oil, fish white rice, tofu, grains, and nuts, a pattern that was similar to that in girls Avoidance (boys): avoidance of fast foods and snacks. <br> Meat(boys): beef and pork, mashed potatoes, and avoiding dark bread | Traditional Western: associated with Chorn's Disease (OR=4.7, 95\% CI 1.6 -14.2) <br> Prudent, : inversely associated with CD in both genders (girls: $\mathrm{OR}=0.3,95 \% \mathrm{CI} 0.1-0.9$; boys: $\mathrm{OR}=0.2,95 \% \mathrm{CI}$ 0.1-0.5) | * |
| $\begin{aligned} & \text { (Kubo et al., } \\ & 2008 \text { ) } \end{aligned}$ | Other(Barrett's Esophagus) | 2. | $2$ <br> Western: French fries, pizza, hamburgers, | health-conscious: inversely associated with Barrett's esophagus (OR : 0.35, 95\% confidence interval: 0.20-0.64) | * |

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|  |  | and tacos), soft drinks, beer/liquor, and coffee and was low in tofu, cooked cereals, fruits, and water. <br> Health-conscious: fruits and vegetables, non-fried fish, and tofu and was low in meat, salty snacks, fried foods, and soft drinks. <br> 3. $12.7 \%$ | Other patterns: Not Statistically Significant |  |
| :---: | :---: | :---: | :---: | :---: |
| (Lutsey et al., 2009) | Other (venous thromboembolism) | 1. 2 <br> 2. Prudent : vegetables, fruit, and poultry Western : processed meat, non-cereal whole grains, and added fats and oils | All the patterns: Not Statistically Significant | Prudent : positively associated with people who were physically active, less likely to be current smokers and took more multivitamin supplements, Western : positively associated with people who had higher BMI, were less physically active, were more likely to smoke and took fewer multivitamin supplements than men |
| $\begin{aligned} & \text { (Fung et al., } \\ & \text { 2004b) } \end{aligned}$ | Other (Stroke) | Same as (Fung et al., 2004a) | Western: associated with an increased relative risk (RR) of 1.58 ( $95 \%$ CI, $1.15-2.15 ; P<0.0002$ for trend) for total strokes and 1.56 ( $95 \%$ CI, $1.05-2.33 ; P<0.02$ for trend) Prudent: associated with decreased relative risk ( RR ) of $0.78(95 \% \mathrm{CI}, 0.61-1.01)$ for total stroke and $0.74(95 \% \mathrm{CI}$, 0.54-1.02) for ischemic stroke. | * |
| (Vujkovic et <br> al., 2009a) | Other (semen quality) | 1. 2 <br> 2. Health Conscious: fruits, vegetables, fish and whole grains Traditional Dutch: meat, potatoes and whole grains and low intakes of beverages and sweets | Health Conscious: inversely correlated with tHcy in blood ( $\mathrm{b}=-0.07, \mathrm{P}<0.02$ ) and seminal plasma $(\mathrm{b}=-1.34, \mathrm{P}<$ 0.02 ) and positively with vitamin B 6 in blood $(\mathrm{b}=0.217, \mathrm{P}$ <0.01) <br> Traditional Dutch: positively correlated with red blood cell folate ( $\mathrm{b}=0.06, \mathrm{P}<0.04$ ) and sperm concentration $(\mathrm{b}=$ $13.25, \mathrm{P}=0.01$ ). | * |
| (Feinstein et <br> al., 2008) | Other (School attentaiment ) | 1. 2 <br> 2. Junk food: high-fat processed foods (sausages, burgers and poultry products), snack foods high in fat and/or sugar (such as crisps, sweets, chocolate, ice lollies and ice creams) fizzy drinks and the number of takeaway meals <br> Health Conscious: vegetarian foods, nuts, salad, rice, pasta, fruit, cheese, fish, cereal | Junk food : negatively associated with the level of school attainment ( $\mathrm{P}<0.05$ ) | * |
| (Rashidkhani et al., 2005) | Other (Renal Cell Carcinoma) | 1. 3 <br> 2. Healthy: vegetables, tomato, fish, fruits, poultry, and whole grains. <br> Western: sweets, processed meat, refined grains, margarine/butter, high fat dairy products, fried potatoes, soft drinks, and | Drinker: associated with decreased risk of Renal Cell Carcinoma risk ( RR comparing the $2^{\text {nd }}$ and $3^{\text {rd }}$ with the first tertile, 0.56 ; $95 \%$ CI, $0.34-0.95$; and $0.72 ; 95 \% \mathrm{CI}, 0.42-$ 1.22 , respectively, P 0.08 by Wald test). <br> Other patterns: Not Statistically Significant | * |

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|  |  | meat <br> Drinker: Alcoholic beverages (wine, liquor, beer) and snacks <br> 3. $25 \%$ |  |  |
| :---: | :---: | :---: | :---: | :---: |
| (Takata et al., 2007) | Other (mammographic density ) | 1. 2 <br> 2. Fat and meat. Vegetables Fruit and milk | Fat and meat: positively associated with mammographic densities than those with lower scores (P or trend=0.21) Vegetables Fruit and milk: weakly inversely associated with mammographic densities only among Japanese women ( P for trend $=0.13$ and 0.03 , respectively). | * |
| (Kontogianni et al., 2009) | Other (lumbar spine bone mineral density) | 1. 10 <br> 2. Component 1: dairy, cereals, red meat, and olive oil consumption <br> Component 2: fish and olive oil and low intake of red meat and products Component 3: poultry and nuts and low intake of red meat and red meat products Component 4,5: alcohol <br> Component 6:legumes <br> Component 7: Sweets <br> Component 8: .fruit drinks <br> Component 9: Coffee <br> Component 10: Soft Drinks <br> 3. $\mathbf{8 0 \%}$ | Component 3: positively associated with lumbar spine bone mineral density (beta coefficient $=0.185, \mathrm{P}<0.017$ and total body bone mineral content (beta coefficient $=0.140 \mathrm{P}$ < 0.048). <br> Other patterns : Not Statistically Significant | * |
| $\begin{aligned} & \text { (Chen et al., } \\ & 2006 \text { ) } \end{aligned}$ | Other (Hypertension) | 1. 3 <br> 2. Balanced: steamed rice, red meat, small fish, fruit, and vegetables Animal protein: fish, eggs, milk, poultry, red meat (beef and mutton), bread, and fruit. Gourd and root vegetable: squashes, pumpkin, sweet potato, radish | Balanced: associated with decreased risk of hypertension ( Adjusted prevalence odds ratios for general hypertension were 1.00 (reference), 0.81 ( $95 \% \mathrm{CI}: 0.79,0.97,0.82$ ( 0.68 , $0.97,0.79(0.66,0.94$, and $0.71(0.59,0.85$ ( P for trend 0.01)) <br> Animal protein: associated with increased risk of hypertension (Prevalence odds ratios for general hypertension were 1.00 (reference), 1.30 (1.01, 1.52., 1.20 (1.01, 1.47, 1.22 ( $1.00,1.44$, and 1.21 ( $1.03,1.49$ (P for trend $=0.23$ )) <br> Other patterns: Not Statistically Significant | Animal protein : positively associated with the prevalence of cigarette smoking and markers of socioeconomic status, including educational attainment, television ownership, and land ownership <br> Gourd and root vegetable : inversely associated with television ownership |
| $\begin{aligned} & \text { (He et al., } \\ & 2009) \end{aligned}$ | Other (higher glucose abnormalities) | 1. 4 <br> 2. Green Water: (like the rice area in the Southeast) <br> Yellow Earth: their food is mainly produced on the dry and hilly land, like the mountain area in the Northwest <br> New Affluence | New Affluence and Yellow Earth: positively associated with higher glucose abnormalities (prevalence ratio 1.22 ( $95 \%$ CI $1.04-1.43$ )) and 2.05 (1.76-2.37). | * |

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|  |  | Western Adopter : Western- oriented food style |  |  |
| :---: | :---: | :---: | :---: | :---: |
| (Mizoue et al., 2006) | Other (Glucose tolerance abnormality) | 1. 3 <br> 2. DFSA (high-dairy, high-fruit and vegetable, high-starch, low-alcohol ): fermented dairy products, milk, confectioneries, bread, fruits, and vegetables, a local alcoholic beverages Animal food: various kinds of animal foods, including red meat, poultry, seafood excluding fish, processed meat and fish products, and fried or broiled foods. <br> Japanese: soybean products, seaweeds, pickles, and green tea, vegetables, and fish <br> 3. $24 \%$ | DFSA (high-dairy, high-fruit and -vegetable, highstarch, low-alcohol ): significantly and inversely associated with a glucose tolerance abnormality (impaired fasting glucose, impaired glucose tolerance, or type 2 diabetes) for the 2 nd , 3 rd , and 4th quartiles were $0.80(95 \% \mathrm{CI}: 0.62-$ 1.04), 0.71 ( $95 \% \mathrm{CI}: 0.54-0.92$ ), and 0.51 ( $95 \% \mathrm{CI}: 0.38-$ 0.67 ), respectively, compared with the lowest quartile). Other patterns : Not Statistically Significant | DFSA (high-dairy, high-fruit and -vegetable, high-starch, low-alcohol ):: positively associated with higher levels of leisure-time physical activity and smaller amounts of alcohol, nonsmoking status Animal food: positively associated with higher BMI and consumed larger amounts of alcohol <br> Japanese: positively associated with higher levels of leisure-time physical activity, consumed greater amounts of alcohol, and had a higher proportion of nonsmokers. |
| $\begin{aligned} & \text { (Tseng et al., } \\ & 2008 \text { ) } \end{aligned}$ | Other (Breast Density ) | 1. 3 <br> 2. Fruit-vegetable-cereals <br> Salad-sauce-pasta/grain: pasta, rice, and such salad and sauce vegetables as mushrooms, garlic, peppers, lettuce, onions, and tomatoes. <br> Meat-starch: French fries, fried chicken and fish, meat, white bread, cheese, eggs, and sweets. | Fruit-vegetable-cereal: inversely associated with breast density among premenopausal women $(b=-0.13, p=0.09$; interaction $\mathrm{p}=0.009$ ) and current smokers, $(\mathrm{b}=-0.30, \mathrm{p}=$ 0.02 ; interaction $\mathrm{p}=0.05$ ), while the salad-saucepasta/grain was inversely associated with breast density among current smokers ( $b=-0.27, p=0.06$; interaction $p=$ 0.006 ). | Fruit-vegetable-cereal and salad-saucepasta/grain : inversely associated with age, college education, living in a large city or a suburb of a large city, with former rather than never smoking, and was positively associated with alcohol intake <br> Meat-starch: positively associated with those who were younger, less well educated, more likely to live in a rural area and to smoke and less likely to use multivitamins or to exercise. |
|  <br> Park, 2007) | Other (Better health status) | 1. 3 <br> 2. Korean healthy: vegetables, kimchi (spicy raw vegetables), seaweeds, beans, fruits, milk and dairy products <br> Animal foods: beef, pork, poultry and fish as well as fast food including hamburgers and pizza. <br> Sweets: high intakes of ice cream, sweet drinks, chocolate, sweet baked goods and sugary foods. <br> 3. $27.4 \%$ | Korean healthy : associated with better health status(as compared with the lowest quintile, the multivariate-adjusted OR of the highest quintile for health status inferior or similar to their peers was $0.59(95 \% \mathrm{CI} 0 \cdot 42,0 \cdot 84)$ <br> Other patterns : Not Statistically Significant | Korean healthy: positively related with people who were more likely to be from the households with higher income and food expenditure, and had mothers with better nutrition attitude. <br> Animal foods: positively related with people who tended to be older and overweight and their households spent more money on buying food. |
| (Ambrosini et al., 2008a) | Other (benign prostatic hyperplasia) | Same as (Ambrosini et al., 2008b) | Vegetable: associated with lower risk of having prostatic hyperplasia (OR $0.78,95 \%$ confidence interval $0.63-0.98$ ). | * |
| (Brantsaeter et al., 2009) | Other ( Pre-eclampsia) | 1. 4. <br> 2. Vegetable: vegetables, cooking oil, olive oil, fruits and berries, rice, and chicken Processed: processed meat products, white bread, French fries, salty snacks, and sugar- | Vegetable: associated with lower risk of Pre-eclampsia (relative risk (OR) for tertile 3 vs. tertile 1:0.72; $95 \% \mathrm{CI}$ : 0.62, 0.85). <br> Processed: associated with increased risk (OR for tertile 3 | Vegetable: positively associated with maternal age, education, and height, supplement use and inversely with BMI, and smoking. <br> Processed: inversely associated with maternal age, |

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|  |  | ( 3. | sweetened drinks and high negative loadings on oily fish, high-fiber breakfast cereals, and lean fish. <br> Potato and Fish Pattern: cooked potatoes, processed fish, lean fish, fish spread and shellfish, and margarine. <br> Cakes and Sweets Pattern: cakes, waffles and pancakes, buns, ice cream, sweet biscuits, sweets, and chocolate. 18\% | vs. tertile 1: $1.21 ; 95 \%$ CI: $1.03,1.42$ ). <br> Other patterns : Not Statistically Significant | supplement use education, and height and positively associated with BMI and smoking, |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Nettleton et al., 2008b) | Other ( Lower spot urine collection) | 1. | 4 <br> Fats and Processed Meat: added fats, processed meat, fried potatoes, and desserts Vegetables and Fish :several vegetable groups, fish, soup, Chinese foods, red meat, poultry, and soy), Beans, Tomatoes, and Refined Grains (beans, tomatoes) Refined grains: dairy foods, avocado/guacamole, and red meat) Whole Grains Fruit (whole grains, fruit, nuts and seeds, green leafy vegetables, and low-fat dairy foods). | Whole grains, fruit, vegetables, and low-fat dairy foods: associated with $20 \%$ lower spot urine collection across quintiles ( P for trend 0.004 ) (renal vascular integrity). | * |
| $\begin{aligned} & \text { (Okubo et al., } \\ & 2007 \text { ) } \end{aligned}$ | Other ( Lower constipation) | 2. | 4 <br> Healthy <br> Japanese traditional <br> Western <br> Coffee and dairy products | Japanese traditional: associated with a significantly lower prevalence of functional constipation. In comparison with the lowest quintile, the multivariate adjusted odds ratio was 0.52 ( $95 \% \mathrm{CI}: 0.41-0.66, \mathrm{p}$ for trend $<0.001$ ). <br> Other patterns : Not Statistically Significant | * |
| $\begin{aligned} & \text { (Okubo et al., } \\ & 2008 \text { ) } \end{aligned}$ | Obesity (BMI) | 2. | 4. <br> Healthy : Green Vegetables, White vegetables, Mushrooms, Seaweeds, Potatoes, Fish and selfish, Fruit, Salted Vegetables <br> Japanese traditional: Soy products, Rice, Miso soup , Fruit and vegetable juices Western : Meats, fats and oils, seasonings , Processed meats, Eggs, butter Coffee and dairy products: Eggs, sugary foods, dairy products | Healthy: significantly associated with a lower risk of BMI>25 (OR of the highest quintile vs. lowest, $0.57 ; 95 \%$ CI: $0.37-0.87$; P for trend 0.05 ). <br> Japanese traditional and Western: significantly associated with an increased risk of BMI>25 (OR: 1.77; $95 \%$ CI: 1.17-2.67; P for trend o0.01 and OR: 1.56; 95\% CI: 1.01-2.40; P for trend: 0.04, respectively). <br> Other patterns: Not Statistically Significant | Healthy: positively associated with smaller number for current smokers and larger number for dietary supplement users and dieters |
| (McNaughton et al., 2007) | Obesity (BMI) | 1. | 3 <br> ethnic foods and alcohol (women) meat, potatoes, and sweet (women) foods; and fruit, vegetables, and | Fruit, vegetables, and dairy: inversely associated with BMI ( $\mathrm{P}<0.004$, waist circumference ( $\mathrm{P}<0.0007$, blood pressure ( $\mathrm{P}: 0.02$., and was positively associated with red cell folate ( $\mathrm{P}<0.03$. | * |

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|  |  | dairy(women) ethnic foods(men) alcohol(men) | Ethnic foods and alcohol: inversely associated with blood pressure (P: 0.008, whereas the meat, potatoes and sweet foods pattern was positively associated with glycated hemoglobin ( $\mathrm{P}<0.01$. <br> Mixed: inversely associated with waist circumference ( $\mathrm{P}<$ 0.02 . and blood pressure ( $\mathrm{P}<0.01$., whereas there were Not Statistically Significant associations with the ethnic foods and alcohol pattern. <br> Other patterns : Not Statistically Significant |  |
| :---: | :---: | :---: | :---: | :---: |
| $\begin{aligned} & \text { (Newby et al., } \\ & \text { 2006ba) } \end{aligned}$ | Obesity ( BMI) | 1. 4 <br> 2. Healthy: Vegetables, fruits and fish Western/Swedish: meat , potatoes refined grains <br> Alcohol, Sweets High-fat dairy and coffee | Healthy: associated with decreased risk in BMI (b: 20.18 $\mathrm{kg} / \mathrm{m} 2$ for a 1 unit increase in SD score, CI: 20.26 to 20.10; $\mathrm{P}, 0.0001$ ), whereas normal weight and overweight women who increased their Healthy pattern score had smaller increases in BMI ( $20.05 \mathrm{~kg} / \mathrm{m} 2$ and $20.11 \mathrm{~kg} / \mathrm{m} 2$, respectively; $\mathrm{P}, 0.05$ for both). <br> Other patterns : Not Statistically Significant | * |
| $\begin{aligned} & \text { (Craig et al., } \\ & 2010 \text { ) } \end{aligned}$ | Obesity | 1. 3 <br> 2. Fruit and vegetables Snacks Fish and sauce <br> 3. $15 \%$ | Snacks: lowest factor score in obese children (P for linear trend $0 \cdot 047$ ). <br> Fish and Sauce: highest factor score in obese children ( P for linear trend 0.023 ). | * |
| $\begin{aligned} & \text { (Paradis et al., } \\ & 2009 \text { ) } \end{aligned}$ | Obesity | 1. 2 <br> 2. Western: refined grains, French fries, red meats, condiments, processed meats and regular soft drinks <br> Prudent: non-hydrogenated fat, vegetables, eggs and fish and seafood. | Western : positively associated with obesity (OR:1.82, 95\% CI 1.16-2.87) <br> Prudent: inversely associated with obesity (OR: 0.62, 95\% CI 0.40-0.96). | Western: positively associated with younger age Prudent : positively associated with older age |
| (Esmaillzadeh \& Azadbakht, 2008b) | Obesity | Same as(Esmaillzadeh \& Azadbakht, 2008a) | Iranian: was associated with increased risk of obesity (subjects in the highest quintile had greater odds of being centrally obese, either before ( $\mathrm{OR}=2.15 ; 95 \% \mathrm{CI} 1 / 4.18-$ 3.90) or after ( $\mathrm{OR}=2.08$; $95 \% \mathrm{CI}: 1.09-3.65$ ) control for confounders) | Same as(Esmaillzadeh \& Azadbakht, 2008a) |
| $\begin{aligned} & \text { (Shi et al., } \\ & 2008 \text { ) } \end{aligned}$ | Obesity | 1. 4. <br> 2. Macho: animal foods and alcohol Traditional: rice and fresh vegetable and inversely on wheat flour sweet tooth :cake, milk, yoghurt and drinks Vegetable- rich: whole grains, fruits, root vegetables, fresh and pickled vegetables, milk, eggs and fish. | Vegetable-rich: independently associated with obesity. Compared with the lowest quartile of vegetable-rich pattern, the highest quartile had higher risk of general obesity (men, prevalence ratio (PR): 1.82, $95 \%$ confidence interval (CI): 1.05-3.14; women, PR: 2.25, $95 \%$ CI: 1.45-3.49). | Vegetable- rich : associated positively with education and negatively with income |
| (Murtaugh et | Obesity | 1. 5 <br> 2. Western : high-fat dairy foods, refined grains and refined grain snacks, gravy and | Western : was associated with higher prevalence of overweight and obesity | * |

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| al., 2007) |  | sauces, fast foods (fries, beef sandwiches, and chicken), bacon and sausage, potatoes, margarine, polyunsaturated fats, high-fat/high-sugar desserts, and red meats Native Hispanic: Mexican cheeses, Mexican soups, Mexican meats, legumes, Mexican tomato-based sauces, and tomato based sauces <br> Prudent: low-fat dairy, whole grain cereals, fruits (canned, dried, and fresh), fruit juices, legumes, vegetables, broth soups, and nuts Mediterranean: liquor, poultry, fish and shellfish, vegetables, salad, Greens and high-fat salad dressings. <br> Dieter: avoiding high-fat dairy products, high-fat salad dressing, cola beverages, and butter, and using low-fat dairy, low-fat margarine, low-fat and fat-free salad dressings, low-fat/high-sugar desserts, diet cola, other diet beverages, and sugar substitutes | Prudent : was associated with a $29 \%$ lower prevalence of overweight <br> Other patterns : Not Statistically Significant |  |
| :---: | :---: | :---: | :---: | :---: |
| $\begin{aligned} & \text { (Pala et al., } \\ & 2006 \text { ) } \end{aligned}$ | Obesity | 1. 3 <br> 2. Prudent: cooked vegetables, pulses, cabbage, seed oil and fish Pasta \& meat: pasta, tomato sauce, red meat, processed meat, bread and wine Olive oil \& salad raw vegetables, olive oil, soup and chicken Sweet \& dairy sugar, cakes, ice cream, coffee and dairy <br> 3. $21 \%$ | Pasta \& meat :associated with increased BMI: $(\mathrm{P}<0.001) .$ | Prudent: positively associated with education Pasta and meat: inversely associated with education |
| (LopezGarcia et al., 2004) | Obesity | 1. 2 . <br> 2. Prudent: vegetables, fruit, legumes, whole grains, fish, and poultry. <br> 3. Western: red meat, processed meat, refined grains, sweets, desserts, French fries, and high-fat dairy products. | Prudent: inversely associated with plasma concentrations of CRP ( $\mathrm{P}<0.02$ ) and E-selectin ( $<0.001$ ) after adjustment for age, body mass index (BMI), physical activity, smoking status, and alcohol consumption. <br> Western : showed a positive relation with CRP ( $\mathrm{P}<$ 0.001 ), interleukin 6 ( $\mathrm{P}<0.006$ ), E-selectin ( $\mathrm{P}<0.001$ ), sICAM-1 ( $\mathrm{P}<0.001$ ), andsVCAM-1 $(\mathrm{P}<0.008)$ after adjustment for all confounders except BMI; Atheroscheloris | Prudent : subjects were more physically active and smoked less |
| (Newby, <br>  <br> Tucker, 2004 | Obesity | 1. 6 <br> 2. Healthy, fiber-rich food pattern (factor <br> 1): reduced-fat dairy products, fruit, and fibre and loaded moderately on fruit juice, non-white bread, nuts and seeds, whole | Factor 1: inversely associated with annual change in BMI(OR=0.51; 95\% CI: $\sim(0.82,0.20) ; P=0.05 ; P$ for trend $=0.01$ )in women and inversely associated with annual change in waist circumference (OR:1.06; 95\% ; P = 0.05; $P$ for trend $=0.04$ ) in both sexes | * |

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|  |  | grains, and beans and legumes <br> Protein and alcohol <br> Sweets <br> Vegetable fats and vegetable <br> Fatty meats <br> Eggs, bread and soup |  |  |
| :---: | :---: | :---: | :---: | :---: |
| $\begin{aligned} & \text { (Togo et al., } \\ & \text { 2004) } \end{aligned}$ | Obesity | 1. 2 for men and 2 for women <br> 2. Green <br> Sweet <br> Traditional <br> Green <br> Sweet-Traditional | Sweet and Sweet-Traditional: inversely associated with baseline BMI. <br> Traditional: inversely associated with subsequent 11- and 5-y BMI change, respectively. <br> Other patterns: Not Statistically Significant | * |
| (Di bello et <br> al. 2008) | Myocardial infraction | 1. 5 <br> 2. Vegetable: fruit, dark yellow vegetables, green leafy vegetables, other vegetables, and polyunsaturated oil and a low intake of palm oil <br> Factors 2 and 3 vegetables and high-fat dairy products including whole milk, ice cream, and cheese <br> Factor 4: palm oil and coffee pattern, was characterized by high intakes of coffee, sugar, and palm oil <br> Factor 5: alcohol, legumes, and polyunsaturated oil pattern | Vegetable: associated with a significantly decreased adjusted risk of $28 \%$ of myocardial infarction. <br> Factor 4: associated with a $38 \%$ increased risk of myocardial infarction <br> Other patterns : Not Statistically Significant | * |
| $\begin{aligned} & \text { (Iqbal et al., } \\ & 2008 \text { ) } \end{aligned}$ | Myocardial infraction | 1. 3 <br> 2. Oriental: tofu and soy and other sauces <br> 3. Western: fried foods, salty snacks, eggs and meat Prudent: fruit and vegetables | Prudent: associated inversely with Acute Myocardial Infarction, (Compared with the first quartile, the adjusted ORs were 0.78 ( $95 \%$ CI 0.69 to 0.88 for the second quartile, $0.66(95 \%$ CI 0.59 to 0.75 for the third, and $0.70(95 \% \mathrm{CI}$ 0.61 to 0.80 ) for the fourth ( P for trend 0.001 ). <br> Western :U-shaped associated with AMI (compared with the first quartile, the adjusted OR for the second quartile was 0.87 [ $95 \%$ CI 0.78 to 0.98 ], whereas it was 1.12 [ $95 \% \mathrm{CI}$ 1.00 to 1.25 ] for the third quartile and 1.35 [ $95 \%$ CI 1.21 to $1.51]$ for the fourth quartile; P for trend 0.001 ) <br> Other patterns: Not Statistically Significant | * |
| (Akesson et al., 2007) | Myocardial Infraction | 1. 4 <br> 2. Healthy: vegetables, fruit, and legumes Western/Swedish: red meat, processed meat, poultry, rice, pasta, eggs, fried potatoes, and fish <br> Alcohol: wine, liquor, beer, and some snacks | Healthy and the alcohol: statistically significantly associated with the risk of primary MI <br> Healthy: associated with $71 \%$ increased risk compared with the highest quintile ( P for trend=.004. In the lowest quintile of the alcohol dietary pattern, the relative risk of MI was 1.64 ( $95 \%$ confidence interval, 1.09-2.47 compared with the | * |

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|  |  | Sweets: sweet baked goods, candy, chocolate, jam, and ice cream | highest quintile ( P for trend=.002). <br> Other patterns: Not Statistically Significant |  |
| :---: | :---: | :---: | :---: | :---: |
| (Ambrosini et al., 2009) | Metabolic syndrome | 1. 2 <br> 2. Western: meat , fast food, potato, soft drinks, cakes, high fat products Healthy: whole grain vegetables and fish <br> 3. $88 \%$ | Western: associated with greater odds for the 'high risk metabolic cluster (highest vs. lowest quartile ( $\mathrm{OR}=2.50$ ,95\% CI: 1.05-5.98;p for trend < 0.02) <br> Other patterns: Not Statistically Significant | * |
| (Noel et al., 2009) | Metabolic Syndrome | 1. 3 . <br> 2. Meat and French fries: meat, processed meat, French fries, pizza and Mexican foods, eggs, alcohol, and other grains and pasta, and low loadings from reduced-fat dairy, fruit and fruit juice, hot and cold cereal, citrus fruit and juice, poultry and vegetables. <br> Traditional: high in beans and legumes, rice, and oil, and low in high-fat dairy, condiments, and nuts and seeds. <br> Sweets: candy, sugar and chocolate candy, soft drinks, sugary beverages, sweet baked goods, dairy desserts, and salty snacks, and low loadings from fish, poultry, vegetables, oils, and soups. | Sweets: After excluding individuals with diabetes, associated with metabolic syndrome (OR: $1.8,95 \% \mathrm{CI}: 1.03$, 3.3). <br> Other patterns: Not Statistically Significant | Sweets: positively associated with younger age, less physical activity, smoking, men, lower vitamin and medication use. |
| (Lutsey, Steffen \& Stevens, 2008) | Metabolic Syndrome | 1. 2 <br> 2. Western: refined grains, processed meat, fried foods, and red meat. <br> Prudent: Cruciferous and carotinoid vegetables, fruit, fish, and poultry. <br> 3. $19.9 \%$ | Western : adversely associated with incident Metabolic Syndrome highest vs. lowest quintile, OR: 1.18 (1.03-1.37) P for trend 0.03 ) <br> Other patterns: Not Statistically Significant | * |
| (Esmaillzadeh et al., 2007) | Metabolic Syndrome | Same as (Esmaillzadeh \& Azadbakht, 2008a) | Healthy: subjects in the highest quintile of pattern scores had a lower odds ratio for the metabolic syndrome (odds ratio: $0.61 ; 95 \% \mathrm{CI}: 0.30,0.79$; P for trend 0.01 . and insulin resistance $(0.51 ; 0.24,0.88 ; \mathrm{P}$ for trend $<0.01$ than did those in the lowest quintile). <br> Western: highest quantile of pattern had greater odds for the metabolic syndrome (1.68; 1.10, 1.95; P for trend $<0.01$ ) and insulin resistance ( $1.26 ; 1.00,1.78$; P for trend $<0.01$ ) <br> Iranian: significantly associated only with abnormal glucose homeostasis (1.19; 1.04, 1.59; P 0.05) | Same as (Esmaillzadeh \& Azadbakht, 2008a) |

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| $\begin{aligned} & \text { (Oddy et al., } \\ & 2009 \text { ) } \end{aligned}$ | Mental Health and Behavioural problems (depression) | 1. <br> 2. <br> 3. | $2$ <br> Healthy: fruits and vegetables and low intakes of crisps and confectionary. <br> Western: meat ,crisps and fast food items and soft drinks $50 \% \text { and } 34 \%$ | Western: associated with higher total ( $\mathrm{b}=2.20,95 \%$ $\mathrm{CI}=1.06,3.35$ ), internalizing (withdrawn/depressed) ( $\mathrm{b}=1.25,95 \% 31 \mathrm{CI}=0.15,2.35$ ) and externalizing (delinquent/aggressive) ( $\mathrm{b}=2.60,95 \% \mathrm{CI}=1.51,3.68$ ) CBCL scores (mental health) <br> Prudent : Not Statistically Significant for the adjusted case | * |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Yannakoulia et al., 2008) | Mental and Health problems (anxiety score) | 1. <br> 2. <br> 3. | 6 <br> Healthful dietary: vegetables, cereals, fish and dairy. <br> Vegetarian: vegetables, fruits, nuts and potatoes <br> Sweets and soft drinks <br> Low-fat/Low-sugar: products such as lowfat dairy products, poultry and light soft drinks <br> Western-type: potatoes, red meat and coffee. <br> Cereals and legumes: cereals and legumes and low alcohol intake <br> 69.3\% | Sweets intake, Meat and products intake, positively associated with anxiety score in females ( $\mathrm{P}<0.05$ ) Cereals and Legumes : negatively associated with anxiety score ( $\mathrm{P}<0.05$ ) | * |
| (Akbaraly et al., 2009) | Mental and Health Problems ( depression) | $\begin{aligned} & \hline 1 . \\ & 2 . \end{aligned}$ | 2 <br> Whole food: vegetables, fruits and fish. Processed food: sweetened desserts, chocolates, fried food, processed meat, pies, refined grains, high-fat dairy products and condiments. | Whole food : associated with lower odds of CES-D depression ( $\mathrm{OR}=0.74,95 \% \mathrm{CI} 0.56-0.99$ ) <br> Processed food: associated with an increased odds of CESD depression ( $\mathrm{OR}=1.58,95 \% \mathrm{CI} 1.11-2.23$ ). | * |
| $\begin{aligned} & \text { (Wiles et al., } \\ & 2009 \text { ) } \end{aligned}$ | Mental and Behavioural Health | $\begin{aligned} & \hline 1 . \\ & 2 . \end{aligned}$ | 3 <br> Junk: high-fat processed foods (burgers, coated poultry) and snack foods high in fat and/or sugar (such as crisps, and chocolate), Health Conscious :rice, pasta, salad and fruit <br> Traditional: diet of meat, potatoes and vegetables | All the patterns: Not Statistically Significant | * |
| (Deshmukh- <br> Taskar et al., 2009) | HDL-Cholesterol | $\begin{aligned} & \hline 1 . \\ & 2 . \end{aligned}$ | 2 <br> Western Dietary Pattern: refined grains, French fries, high-fat dairy products, dishes with cheese, red meats, processed meats, eggs, snacks, sweets and desserts, sweetened beverages and condiments Prudent Dietary Pattern: grains, legumes, vegetables (i.e. cruciferous, other leafy and dark-yellow vegetables), tomatoes, fruits, | Prudent:: positively associated with Insulin sensitivity ( $\mathrm{P}<$ 0.005) <br> Western: inversely associated with Serum HDL cholesterol ( $\mathrm{P}<0.005$ ) | * |

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\begin{tabular}{|c|c|c|c|c|c|}
\hline \& \& \& 100\% fruit juices, low-fat dairy products, poultry, clear soups and low-fat salad dressings)
\[
31 \% .
\] \& \& \\
\hline \begin{tabular}{l}
(Hamer \& \\
Mishra, 2010)
\end{tabular} \& HDL-cholesterol \& 1. \& \begin{tabular}{l}
4. \\
Fast food: pasta , chicken, burger , kebab, crisps, chocolate, tea, coffee Health aware: wholemeal bread ,oily fish ,vegetables, yogurt \\
Traditional: white bread ,eggs , bacon , sausages, beer Sweet: breakfast cereals, biscuits, buns, tea, coffee, beer 16.5\%
\end{tabular} \& Health aware: inversely associated with concentrations of homocysteine (mean \(=-2.43,95 \% \mathrm{CI}:-3.41,-1.45, \mathrm{P}\) trend 0.01 ) and, and positively with HDL-cholesterol ( highest vs. lowest tertile mean \(=0.08(95 \% \mathrm{CI}, 0.03,0.12\); P trend \(=0.09)\) Other patterns : Not Statistically Significant \& Fast food : positively associated with younger age, smoking, employment and single life \\
\hline (Panagiotakos et al., 2007b) \& HDL- Cholesterol \& \& \begin{tabular}{l}
6 \\
Component \(\mathbf{1 / h e a l t h f u l ~ : ~ l o w - f a t ~ p r o d u c t s ~}\) such as fish, vegetables, legumes, cereals, and fruits; \\
Component 2/ high glycemic index and Component \(\mathbf{3 / h i g h}-\) fat : red or white meat and meat products, and potatoes Component 4/pasta: pattern that included consumption of pasta Component 5/dairy products and eggs Component 6/sweets Component 7/alcohol
\end{tabular} \& \begin{tabular}{l}
Component 1 : inversely associated with waist circumference, systolic blood pressure, triglycerides, positively associated with high-density lipoprotein cholesterol levels, and inversely with the likelihood of the metabolic syndrome (odds ratio [OR] \(0.87,95 \%\) confidence interval [CI] 0.79 to 0.97 , \\
Components 2 and 6 : positively correlated with the previous indexes, and the likelihood of having the metabolic syndrome (OR \(1.13,95 \%\) CI 1.05 to 1.21 and OR 1.26, 95\% CI 1.21 to 1.33 ) \\
Other patterns : Not Statistically Significant
\end{tabular} \& \begin{tabular}{l}
Component 1 was inversely associated with age and positively associated with male sex and physical activity status. \\
Component 2 was positively associated with male sex. \\
Component 3 was positively associated with age income and physical activity level. \\
Component 4 was positively associated with age and smoking habits and \\
Component 5 was positively associated with male sex, age.
\end{tabular} \\
\hline (Sadakane et al., 2008) \& HDL Cholesterol \& 2.

3. \& \begin{tabular}{l}
3 <br>
Vegetable: vegetables, potatoes, soybeans products tofu and fermented soybeans, fruits, sea weeds, citrus, beans, and dried fish <br>
Meat: processed meats, beef, pork, poultry, steamed fish paste, high-fat products, and butter <br>
Western: breads, butter, and yoghurt, and lower intakes of rice, salty products, and miso soup 28.5\%

 \& Vegetable: associated with higher HDL cholesterol. Meat: associated with higher total and HDL cholesterol. Western : associated with higher total, HDL, and LDL cholesterol \& 

Vegetable: positively associated with people who were older, married, and less likely to smoke. <br>
Meat: positively associated with people who were younger, married, highly educated, and likely to smoke and drink alcohol. <br>
Western: positively associated with people who were younger, highly educated, less likely to smoke and drink alcohol, but were physically inactive.
\end{tabular} <br>

\hline \[
$$
\begin{aligned}
& \text { (Zhang et al., } \\
& 2006 \text { ) }
\end{aligned}
$$

\] \& Diabetes \& \& | 2 |
| :--- |
| Prudent: fruits, vegetables, whole grains, fish, and poultry Western: meat products (red meat and | \& | Western: increased risk of gestational diabetes mellitus (RR: 1.63 ( $95 \%$ CI 1.20-2.21,p for trend $=0.001$ ) |
| :--- |
| Prudent : RR: $1.39(95 \%$ CI $1.08-1.80$, p for trend $=0.018$. | \& Western : positively associated with people who tended to smoke more, consume less fibre, and engage in less physical activity <br>

\hline
\end{tabular}

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|  |  | processed meat), refined grains and high-fat dairy <br> 3. $15.2 \%$ |  | Prudent : positively associated with people who tended to smoke less, consume more fibre, and to be more physically active |
| :---: | :---: | :---: | :---: | :---: |
| $\begin{aligned} & \text { (Cai et al., } \\ & 2007 \text { ) } \end{aligned}$ | CVD, T2D, Stroke, Diabetes | 1. 3 <br> 2. Vegetable-rich: green beans and yard long beans <br> 3. Fruit-rich: fruits Meat-Rich: meat, poultry, and animal organs | Fruit-rich:decreased risk of all causes of death( HR: 0.94 (95\% CI: 0.89-0.98 ) <br> Meat-rich : associated with increased risk of diabetes (HR $1.18,95 \% \text { CI: } 0.98-1.42)$ | * |
| (Nettleton et al., 2009) | CVD | 1. 4 <br> 2. Fats and Processed Meat: added fats, processed meat, fried potatoes, and desserts Vegetables and Fish :several vegetable groups, fish, soup, Chinese foods, red meat, poultry, and soy, Beans, Tomatoes, and Refined Grains (beans, tomatoes) Refined grains: high-fat dairy foods, avocado/guacamole, red meat, and whole grains <br> Fruit: whole grains, fruit, nuts and seeds, green leafy vegetables, and low-fat dairy foods. | Fats and Processed Meat : associated with a greater risk of incident CVD (hazard ratio quintile 5 compared with quintile 1: $1.82 ; 95 \% \mathrm{CI}: 0.99,3.3$ ) <br> Fruit: associated with a lower risk of CVD $(0.54 ; 95 \% \mathrm{CI}$ : $0.33,0.91$ ) <br> Other patterns : Not Statistically Significant | * |
| (Esmaillzadeh \& Azadbakht, 2008a) | CVD | 1. 3 <br> 2. Healthy: fruits, vegetables, tomatoes, poultry, legumes, cruciferous and green leafy vegetables, tea, fruit juices, and whole grains <br> Western: refined grains, red meat, butter, processed meat, high-fat dairy products, sweets and desserts, pizza, potatoes, eggs, hydrogenated fats, and soft drinks. Low in other vegetables and low fat dairy products Iranian: refined grains, potato, tea, whole grains, hydrogenated fats, legumes, and broth <br> 3. $24 \%$ | Healthy: decreased risk of dyslipidemia (odds ratio (OR), $0.36 ; 95 \% \mathrm{CI}, 0.19-0.53$ ), hypertension (OR, $0.33 ; 95 \% \mathrm{CI}$, $0.17-0.60)$, at least 1 (OR, $0.30 ; 95 \% \mathrm{CI}, 0.18-0.58$, and at least 2 risk factors (OR, $0.39 ; 95 \% \mathrm{CI}, 0.20-0.77$ ) compared with the lowest quintile. <br> Western: increased cardiovascular risk factors (OR, 2.593.11; $\mathrm{P}<0.05$ ). <br> Iranian: significantly associated with dyslipidemia (OR, 1.73; $95 \% \mathrm{CI}, 1.02-2.99$ and at least 1 risk factor (OR, 1.89; 95\% CI, 1.05-3.20). | Healthy: positively associated with more physical activity and greater fiber intake but lower energy and cholesterol intakes. <br> Western positively associated with less physical activity and lower fiber intake but greater energy and cholesterol intakes. <br> Iranian: positively associated with age, physical activity, lower energy intake |
| (Heidemann et al., 2008) | CVD | 1. 2 <br> 2. Prudent: vegetables, fruit, legumes, fish, poultry, and whole grains. <br> Western: red meat, processed meat, refined grains, french fries, and sweets/desserts | Prudent: associated with a $28 \%$ lower risk of cardiovascular mortality ( $95 \%$ CI: 13-40) and a $17 \%$ lower risk of allcause mortality ( $95 \%$ CI: $10-24$ ) <br> Western : associated with a higher risk of mortality from cardiovascular disease $22 \%$ ( $95 \%$ CI: $1-48$ ), cancer ( $16 \%$; $95 \%$ CI, 3-30), and all causes $21 \%$; $95 \%$ CI: 12-32) | Prudent: associated with people who were slightly older, exercised more, were less likely to be smokers, were more likely to use hormone replacement therapy and multivitamin supplements, and had a more advantageous nutrient profile than those with lower sores for this pattern. |

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|  |  |  |  | Western: associated with people who were younger, less physically active, were more likely to smoke, were less likely to use hormone replacement therapy and multivitamin supplements, and had a more unfavorable nutrient profile than those who scored low on this pattern. |
| :---: | :---: | :---: | :---: | :---: |
| (Panagiotakos et al., 2007a) | CVD | 1. 5 <br> 2. Component 1: low-fat products, like fish, vegetables, legumes, greens, and salads Component 2: red or white meat and meat products, pasta, and potatoes <br> Component 3: cereals and sweets <br> Component 4: dairy products and fruits Component 5: alcoholic beverages <br> 3. $56.3 \%$ | Component 1, Component 3, and component 5 were associated with lower likelihood of having increased burden of CVD ( $\mathrm{P}<0.01$ ), irrespective of various potential confounders. <br> Other patterns : Not Statistically Significant | Component 1: inversely associated with age and number of cigarettes smoked and positively associated with physical activity status. <br> Component 2: positively associated with age, male sex and inversely associated with income and physical activity level. |
| (Shimazu et <br> al., 2007) | CVD | 1. 3 <br> 2. Japanese: soybean products, fish, seaweeds, vegetables, fruits and green tea Animal food: animal-derived foods (beef, pork, ham, sausage, chicken, liver and butter), coffee and alcoholic beverages. high-dairy, high-fruit-and-vegetable, and low-alcohol: dairy products milk and yoghurt), margarine, fruits and vegetables (carrot, pumpkin and tomato <br> 3. $26.2 \%$ | Japanese: associated with a lower risk of CVD mortality (hazard ratio of the highest quartile vs. the lowest, 0.73 ; $95 \%$ confidence interval: $0.59-0.90$; P for trend: 0.003... <br> Animal food: associated with an increased risk of CVD. Other patterns : Not Statistically Significant | * |
| $\begin{aligned} & \text { (Tseng et al., } \\ & 2004 \text { ) } \end{aligned}$ | Cancer(prostate) | 1. 3 <br> 2. Vegetable fruit: vegetables, fruits, fish, and shellfish <br> Red meat-starch: red meats, potatoes cheese, salty snacks, and desserts Southern pattern: cornbread, grits, sweet potatoes, okra, beans, and rice. | Southern : decreased risk of prostate cancer (3rd versus 1st tertile relative risk, $0.6 ; 95 \%$ confidence interval, 0.4 1.1 ; trend $\mathrm{P}<0.08$ ) <br> Other patterns: Not Statistically Significant | * |
| $\begin{aligned} & \text { (Chang et al., } \\ & 2008 \text { ) } \end{aligned}$ | Cancer(Ovarian) | 1. 5 <br> 2. Plant-based : Vegetables and Fruits High- Protein / High-Fat: processed meat, butter, ice-cream, cheese and potatoes High-Carbohydrate: pizza, spaghetti , cheese <br> Ethnic: Beans, lentils, rice, potatoes Salad and wine <br> 3. $35.0 \%$ | plant-based: increased relative risk of ovarian cancer was 1.65 ( $95 \%$ confidence interval: $1.07-2.54$; P for trend $=0.03$ ) Other patterns: Not Statistically Significant | * |

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| (De Stefani et <br> al., 2008b) | Cancer(Lung) | 1. 3 <br> 2. Antioxidants Carbohydrates High Meat | Antioxidants: inversely associated with lung cancer risk (OR 0.69, 0.51-0.96 <br> High-meat: associated with a strong increase in risk (OR 2.90, $95 \%$ CI 1.91-4.40). <br> Other patterns : Not Statistically Significant | High meat: positively associated with younger agerural living, inversely correlated with education and associated with smoking intensity and duration and positively correlated with mate drinking Carbohydrates: inversely associated with age. Antioxidants: associated with socio demographic variables, smoking and alcohol drinking. |
| :---: | :---: | :---: | :---: | :---: |
| (Campbell, <br>  <br> Kreiger, <br> 2008) | Cancer(gastric) | 1. 2 <br> 2. Prudent: vegetables, fruits, and fish Western: soft drinks, French fries, white bread, hamburger, eggs, bacon, doughnuts, and hot dogs <br> 3. $23 \%$ | Prudent: associated with decreased risk of gastric cancer in women (odds ratio (OR): $0.58,95 \%$ confidence interval (CI): 0.37, 0.92) <br> Western: associated with increased risk of gastric cancer in women (OR: 1.86, $95 \%$ CI: 1.20, 2.89 and men (OR: 1.44, 95\% CI: 1.03, 2.02.) | * |
| (De Stefani et al., 2004) | Cancer(gastric) | 1. 3 <br> 2. Starchy: total grains and tubers. <br> Healthy: white meat, dairy foods, desserts, raw vegetables, and fruits. <br> Mixed pattern: red meat, processed meat, eggs, and pulses. | Starchy : increased risk of gastric cancer (odds ratio (OR): 4.1, $95 \%$ confidence interval CI: 2.6-6.6) | * |
| (Reedy et al., 2010) | Cancer(Colorectal) | 1. 3 <br> 2. fruits and vegetables fat reduced: meat and potatoes | Meat and potatoes: associated with increased risk of colorectal cancer risk. <br> Other patterns : Not Statistically Significant | * |
| $\begin{aligned} & \text { (Butler et al., } \\ & \text { 2008) } \end{aligned}$ | Cancer(Colorectal) | 1. 2 <br> 2. Vegetable-fruit-soy: vegetable, fruit, and soyfood intake; of the 32 foods included in the pattern, 23 were vegetables, five were soyfood items, and four were fruit items. Meat-dim sum: chicken, pork, fish, rice and noodle dishes, and preserved foods. | All the patterns: Not Statistically Significant | meat-dim sum : positively associated with being male, have higher education, report any weekly physical activity, be a heavy smoker, drink alcohol, and consume more saturated fat, compared to individuals in the first quartile <br> Vegetable-fruit-soy positively associated with physical activity and education level and inversely with heavy smoking. |
| (Edefonti et <br> al., 2008) | Cancer(Breast) | 1. 4 <br> 2. Animal products: animal protein and animal fat, calcium, cholesterol, saturated fatty acids, riboflavin, zinc and phosphorus Vitamins and Fiber: vitamin $C$ and total fiber, total folate, potassium, b-carotene equivalent, soluble carbohydrates and vitamin <br> Unsaturated fats: vegetable fat and vitamin E, monounsaturated and | Animal products pattern and the Unsaturated fats pattern :negatively associated with breast cancer ( $\mathrm{OR}=0.74,95 \%$ CI: $0.61-0.91$ ) and ( $\mathrm{OR}=0.83,95 \%$ CI: $0.68-1.00$ ), <br> Starch-rich associated with breast cancer ( $\mathrm{OR}=1.34,95 \%$ CI: 1.10-1.65) <br> Vitamins and fiber : inversely associated with ovarian cancer ( $\mathrm{OR}=0.77,95 \% \mathrm{CI}: 0.61-0.9$ ) <br> Starch-rich pattern: positively associated with ovarian cancer (OR = $1.85,95 \%$ CI: 1.37-2.48) | * |

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|  |  | polyunsaturated fatty acids <br> Starch-rich: starch, vegetable protein and sodium. |  |  |
| :---: | :---: | :---: | :---: | :---: |
| (Mannisto et al., 2005) | Cancer(Breast) | 1. 3 . <br> 2. Vegetables: vegetables, legumes, fruit, pasta, fish and oil Pork, Processed meat, Potatoes - PPP: pork, beef, processed meats, potatoes, rice, poultry, liver, butter/low-fat margarine, pasta and coffee <br> 3. $23.2 \%, 29.0 \%$ and $21.8 \%$ | All patterns : Not Statistically Significant | * |
| (De Stefani et <br> al., 2008a) | Cancer(Bladder) | 1. 3 <br> 2. Sweet beverages: coffee, tea, and added sugar. <br> Prudent: fresh vegetables, cooked vegetables, and fruits. <br> Western: red meat, fried eggs, potatoes, and red wine <br> 3. $25.1 \%$ | Sweet beverages: associated with risk of bladder cancer (OR 3.27, 95\% CI: 1.96-5.45) <br> Prudent pattern: associated with risk of bladder cancer (OR 2.35, 95\% CI: 1.42-3.89. <br> Other patterns : Not Statistically Significant | * |
| $\begin{aligned} & \text { (Cottet et al., } \\ & 2005 \text { ) } \end{aligned}$ | Cancer (reduced adenoma) | 1. 3 <br> 2. Mediterranean: olive oil, fresh fruit, vegetables, legumes, lean meat and fresh fish. <br> Western: potatoes, fats, delicatessen products, high-fat meat, beer, rice and pasta, refined bread and cereals, nuts, sodas. <br> Snacks: high fat delicatessen, high-fat cheese, desserts and sweets, beer, soda and mineral water <br> 3. $21.9 \%$ | Mediterranean: associated with reduced adenoma recurrence (second tertile: adjusted odds ratio $(\mathrm{OR})=0.50$, $95 \%$ confidence interval $(C I)=0.18-1.42$; third tertile: adjusted $\mathrm{OR}=0.30,95 \% \mathrm{CI}=0.09-0.98 ; \mathrm{P}$ for linear trend $=$ 0.04). <br> Other patterns: Not Statistically Significant | * |
| (Jackson et <br> al., 2009) | Cancer (prostate) | 1. 4 . <br> 2. Healthy : vegetables, fruits and peas and beans <br> Carbohydrate white bread and refined cereals <br> Sugary foods and Sweet Baked Products Organ meat and fast food: high fat dessert, organ meat, fast food and salty snacks <br> 3. $24.5 \%$ | All the patterns: Not Statistically Significant | Carbohydrate : positively associated with lower use of vitamins, less obesity Sugary foods and sweet baked goods positively associated with smoking and were less likely to have tertiary education. Men with high intakes of the Organ meat and fast food : positively associated with obesity |
| $\begin{aligned} & \text { (Muller et al., } \\ & \text { 2009) } \end{aligned}$ | Cancer (Prostate) | 1. 4 <br> 2. Mediterranean: meats, vegetables, and fruits, and avoidance of cakes and sweet biscuits | All the patterns: Not Statistically Significant | * |

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$\left.\begin{array}{|l|l|ll|l|l|}\hline & & & \begin{array}{l}\text { Vegetable: vegetables } \\ \text { Meat \& Potatoes: meats and potato cooked } \\ \text { in fat }\end{array} \\ \text { Fruit \& Salad : high intake of salad greens } \\ \text { and fruit. }\end{array}\right]$

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| (Balder, Goldbohm \& van den Brandt, 2005) | Cancer (Lung) | $\begin{aligned} & 1 . \\ & 2 . \end{aligned}$ <br> 3. | 5 <br> Salad vegetables: vegetable items, several fruit items, pasta, rice, poultry, fish, and oil. Cooked vegetables: cooked leaf vegetables, cabbages, legumes, and carrots. <br> Pork, processed meat, and potatoes: also coffee and low-fat margarine, Sweet foods: cakes and cookies, sweet sandwich spread, sweets and candies, and (straw)berries <br> White/brown bread : brown/ wholemeal bread types and apples and pears 23\% | Salad vegetables: was associated with decreased risk of lung cancer (rate ratios (RR) Q5, 0.75 ; $95 \%$ confidence interval (CI): 0.55-1.01], after multivariate adjustment. <br> Sweet foods: was inversely associated with lung cancer risk (RR Q5, 0.62; 95\% CI, 0.43-0.89. <br> Other patterns: Not Statistically Significant | * |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\begin{aligned} & \text { (Sant et al. } \\ & 2007 \text { ) } \end{aligned}$ | Cancer (HER-2-positive cancers) | $\begin{aligned} & 1 . \\ & 2 . \end{aligned}$ $3 .$ | 4 <br> Salad Vegetables: raw vegetables and olive oil <br> Western: potatoes, ravioli, red and processed meat, eggs, butter, seed oil (as added fat) and cakes. <br> Canteen: pasta, tomato sauce, olive oil and wine. <br> Prudent: cooked vegetables, rice, poultry, fish and low consumption of alcohol. 21\% | Salad vegetables: decreased risk wit HER-2- positive cancers (RR $50.25,95 \%$ CI: $0.10-0.64$, for the highest tertile; p trend $<0.001$, much stronger than for HER-2negative cancers (heterogeneity < 0.039). <br> Other patterns : Not Statistically Significant | * |
| (Bastos et al., 2010) | Cancer (gastric) | $\begin{aligned} & 1 . \\ & 2 . \end{aligned}$ | 3 <br> Dietary pattern I: fruits and dairy products, and low consumption of alcoholic beverages. <br> Dietary pattern II: fruit, salads, vegetables, dairy products, fish and meat. <br> Dietary pattern III: most food groups and low vegetable soup intake. | Dietary pattern II (low consumption of fruit, salads, vegetables, dairy products, fish and meat): associated with higher gastric cancer compared to Pattern I ( $\mathrm{OR}=1.68$, 95\%CI: 1.31-2.14). Similar associations for cardiac and noncardiac gastric cancer. <br> Other patterns : Not Statistically Significant | * |
| (Bertuccio et <br> al., 2009) | Cancer (gastric) | 1. | 4. <br> Animal products: animal protein, riboflavin, cholesterol, phosphorus, calcium, and zinc. <br> vitamins and fiber: vitamin $C$, total fiber, potassium, total folate, $\beta$-carotene equivalents, and soluble carbohydrates vegetable: polyunsaturated fatty acids, vitamin E , monounsaturated fatty acids, linoleic acid, and linolenic acid <br> Starch-rich: starch, vegetable protein, and | Animal products: associated positively with gastric cancer ( $\mathrm{OR}=2.13$; 95\% CI, 1.34-3.40) <br> Starch-rich: associated positively with gastric cancer (OR=1.67; 95\% CI, 1.01-2.77) <br> Vitamins and fiber: associated inversely with gastric cancer ( $\mathrm{OR}=0.60$; 95\% CI, 0.37-0.99 ) | * |

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|  |  | sodium. |  |  |
| :---: | :---: | :---: | :---: | :---: |
| (Kim et al., 2004) | Cancer (gastric) | 1. 3 <br> 2. Prudent: vegetables, fruits, soy products, seaweeds, mushroom, milk, beans and yogurt <br> Traditional: pickled vegetables, salted fish and roe, fish, rice and miso soup, alcoholic beverages (sake, shochu and beer) for men and thus was called the traditional dietary pattern. <br> Western: meat, poultry, cheese, bread and butter | Prudent: inverse associated between the healthy pattern and gastric cancer risk in women (rate ratio for highest quartile (RR) : $0.56 ; 95 \% \mathrm{CI}: 0.32-0.96 ; p$ for trend $<0.03$ ) Traditional: significantly associated with the increased risk of gastric cancer in both genders (for men, RR: 2.88, 95\% CI: 1.76-4.72; for women, RR: $2.40,95 \%$ CI: 1.32-4.35). Other patterns: Not Statistically Significant | Prudent :positively related with subjects who were more physically active and smoked less |
| $\begin{aligned} & \text { (Flood et al., } \\ & 2008 \text { ) } \end{aligned}$ | Cancer (Colorectal) | 1. 3 <br> 2. First factor (men/women): fruit and vegetables <br> Second factor(men): fat-reduced foods, diet foods, and lean meats Second factor (women): high-fat foods, red meats, and potatoes. <br> Third factor(men): processed foods (sausages and French fries) and sweets Third factor(women): chicken, milk, mayo <br> 3. $35.1 \%$ and $34.2 \%$ | Fruit and vegetables: associated with decreased risk of colorectal cancer (relative risk (RR) for quintile Q 5 versus Q1: $0.81 ; 95 \% \mathrm{CI}: 0.70,0.93$; P for trend 0.004 ). <br> Red meat and potatoes: associated with increased risk: men (RR: 1.17; 95\% CI: 1.02, 1.35; P for trend 0.14 and women (RR: $1.48 ; 95 \%$ CI: $1.20,1.83 ; \mathrm{P}$ for trend 0.0002 ) | Fruit and vegetable: associated with lower BMI, physically activity, education, non-smoking, and less alcohol. <br> Red meat and potatoes: associated with higher BMI, increased energy intake, decreased physical activity, a lower likelihood of being a college graduate, and increased smoking for both men and women. |
| (Kesse, <br> Clavel- <br>  <br> Boutron- <br> Ruault, 2006) | Cancer (Colorectal) | 1. 4 <br> 2. Healthy: raw and cooked vegetables, legumes, fruit, yogurt, fresh cheese, breakfast cereals, sea products, eggs, and vegetable oils (olive oil and others) and by a low consumption of sweets Western: potatoes, pizza and pie, sandwiches, legumes, sweets, cakes, cheese, bread, rice, pasta, processed meat, eggs, and butter <br> Drinker: sandwiches, snacks, coffee, processed meat, sea products, wine, and other alcoholic beverages, as well as by a low consumption of soup and fruit, strongly associated with ethanol intake. <br> Meat eaters: potatoes, legumes, coffee, meat, poultry, vegetable oils (except olive oil), and margarine and negatively | Western: positively associated with adenoma (For quartile 4 versus quartile 1 , an increased risk of adenoma was observed with high scores of the Western pattern (multivariate relative risk (RR): $1.39,95 \%$ confidence interval: 1.00, 1.94; ptrend $<0.03$ ) and the drinker pattern (RR: 1.42, $95 \%$ confidence interval: $1.10,1.83$; ptrend $<$ $0.01)$. <br> Other patterns : Not Statistically Significant | * |

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|  |  | associated with tea, olive oil, and breakfast cereals. |  |  |
| :---: | :---: | :---: | :---: | :---: |
| (Mizoue et al., 2005) | Cancer (Colorectal) | 1. 3 <br> 2. DFSA: A high-dairy, high-fruit and vegetable, high starch, low-alcohol Animal food Japanese | DFSA: Dairy products and fruits and vegetables with low alcohol consumption: associated with decreased risk of colorectal adenomas. <br> Other patterns: Not Statistically Significant | * |
| (Meyerhardt et al., 2007) | Cancer (Colon) | 1. 2 <br> 2. Prudent: fruits, vegetables, whole grains, legumes, poultry, and fish Western: refined grains, processed and red meats, desserts, high-fat dairy products, and French fries | Western: associated with a significantly worse disease-free survival (colon cancer recurrences or death). Compared with patients in the lowest quintile of Western dietary pattern, those in the highest quintile experienced an adjusted hazard ratio (AHR) for disease-free survival of 3.25 ( $95 \%$ CI: 2.045.19; $P$ for trend< .001 ). Also associated with a similar detriment in recurrence free survival (AHR, $2.85 ; 95 \% \mathrm{CI}$, 1.75-4.63) and overall survival (AHR, $2.32 ; 95 \% \mathrm{CI}, 1.36-$ 3.96)), comparing highest to lowest quintiles (both with $P$ for trend <.001). <br> Other patterns: Not Statistically Significant | Prudent: positively associated with physical activity lower body mass no smoking cigarettes <br> Western: positively correlated with white men, and past or current smokers. |
| $\begin{aligned} & \text { (Wu et al., } \\ & \text { 2004a) } \end{aligned}$ | Cancer (Colon) | 1. 2 <br> 2. Prudent : fruits, vegetables, poultry Western : meat,eggs, high-sugar foods | Prudent : weakly and Not Statistically Significantly associated with decreased risk of colon cancer or distal colon adenoma (highest versus lowest quintile: colon cancer: multivariate adjusted relative risk $(R R)=0.84(95 \%$ confidence interval (CI) (0.64-1.10); p trend 0.37) distal adenoma: multivariate odds ratio (OR):0.88 (95\% CI: 0.731.08; p trend= $=0.12$ <br> Western: positively associated with colon cancer and distal adenoma (colon cancer: RR:1.27 (95\% CI:0.96-1.69, p trend $=0.05$; distal adenoma: $\mathrm{OR}=1.28$ ( $95 \% \mathrm{CI}: 1.05-1.56$, p trend 0.01 ) | * |
| $\begin{aligned} & \hline \text { (Cottet et al., } \\ & \text { 2009) } \end{aligned}$ | Cancer (Breast) | 1. 2. <br> 2. Western/alcohol: processed meat and meat products (ham, offal), French fries, appetizers, sandwiches, rice/pasta, potatoes, pulses, pizza/pies, canned fish, eggs, crustaceans, alcoholic beverages, cakes, mayonnaise, and butter/cream. <br> Healthy/Mediterranean: vegetables and fruits, fish and crustaceans, olives, and sunflower oil. traits (fish, fruits, vegetables, olive oil, essentially vegetables, fruits, seafood, olive oil, and sunflower oil) | Western/alcohol: positively associated with breast cancer risk (hazard ratio $=1.20,95 \% \mathrm{CI}: 1.03,1.38 ; \mathrm{P}<0.007$ for linear trend). <br> Healthy/Mediterranean: negatively associated with breast cancer risk (hazard ratio $=0.85,95 \%$ CI: $0.75,0.95 ; \mathrm{P}<$ 0.003 for linear trend), | Western/alcohol: positively associated with younger age, decreasing prevalence of null parity, decreasing duration of breastfeeding, increasing prevalence of overweight, greater height, a higher proportion of relatives with a history of breast cancer, a higher proportion of oral contraceptive use, biennial Pap smears and a higher prevalence of current smoking. <br> Healthy/Mediterranean : positively associated with older age, higher education, a higher prevalence of overweight, a higher proportion of personal history of benign breast disease, increasing use of |

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|  |  |  |  | menopausal hormone therapy, increasing duration of breastfeeding, an increasing proportion of annual Pap smears, higher levels of physical activity, and an increasing proportion of former smokers. |
| :---: | :---: | :---: | :---: | :---: |
| $\begin{aligned} & \text { (Wu et al., } \\ & 2009 \text { ) } \end{aligned}$ | Cancer (breast) | 1. 3 <br> 2. Western meat/starch: pasta with meat, beef taco, beef burrito, pizza, meatballs, hamburger, fried potatoes, baked potatoes, mashed potatoes, pancake, bagels Ethnic meat/starch: pork/fish soups, liver, pork spareribs, salted/dried fish, fried shellfish, chicken wings, rice, fried/Spanish rice, fried noodles, fried dim sum, and other foods <br> Vegetables/soy: green beans/peas, carrots, cabbage, bean sprouts, green peppers, bok choy, fresh tofu, fresh soybeans, soybean milk, and other items | Western: increased risk of breast cancer (OR: 2.19; 95\% CI: 1.40, 3.42; P for trend $<0.0005$ ). <br> Other patterns: Not Statistically Significant | * |
| (Murtaugh et al., 2008) | Cancer (Breast) | 1. 5 <br> 2. Western: high-fat dairy foods, refined grains, gravy and sauces, fast foods, red and processed meats, potatoes, margarine, polyunsaturated fats, and high-fat and highsugar desserts. <br> Native Mexican: Mexican cheeses, soups, meat dishes, legumes, and tomato-based sauces <br> Prudent: low-fat dairy, whole grains, fruit and fruit juice, legumes, vegetables, and soups <br> Mediterranean: liquor, poultry, seafood, vegetables, salad greens, and high-fat salad dressings <br> Diet: high-fat dairy products and salad dressing, cola beverages, and butter and with using low-fat dairy, margarine, and salad dressings, as well as low-fat highsugar desserts, diet beverages, and sugar substitutes | Western: increased risk of breast Cancer (odds ratio for highest versus lowest quartile: $1.32 ; 95 \%$ CI: $1.04,168 ; \mathrm{P}$ for trend 0.01) <br> Prudent: increased risk of breast cancer (OR: 1.42; 1.14, 1.77; P for trend $<0.01$ ) <br> Native Mexican : increased risk of breast cancer (OR 0.68; $0.55,0.85$; P for trend $<0.01$ ) <br> Mediterranean : increased risk of breast Ccancer (0.76; $0.63,0.92$; P for trend $<0.01$ ) <br> Other patterns: Not Statistically Significant | * |
| $\begin{aligned} & \text { (Cui et al., } \\ & \text { 2007) } \end{aligned}$ | Cancer (Breast) | 1. 2 <br> 2. Vegetable-soy: variety of different vegetables, soy-based products, and freshwater fish. | Meat-sweet: associated with increased breast cancer risk in postmenopausal Chinese women. (4 $4^{\text {th }}$ versus 1st quartile: odds ratio, 1.3; $95 \%$ confidence interval, 1.0-1.7; Ptrend $=$ | Vegetable-soy: positively associated with people who were more likely to be physically active. Meat-sweet: positively associated with people who |

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|  |  | Meat-sweet: various meats, primarily pork but also poultry, organ meats, beef and lamb, and shrimp, saltwater fish, and shellfish, as well as candy, dessert, bread, and milk | 0.03 ., but only in postmenopausal women, specifically among those with estrogen receptor-positive tumors (4th versus 1st quartile: odds ratio, $1.9 ; 95 \%$ confidence interval, 1.1-3.3; Ptrend $=0.03$ ). <br> Other patterns: Not Statistically Significant | were younger, better educated, had later menopausal age, and were less likely to have had their first live birth before age 25 years. |
| :---: | :---: | :---: | :---: | :---: |
| $\begin{aligned} & \text { (Hirose et al., } \\ & \text { 2007) } \end{aligned}$ | Cancer (Breast) | 1. 4 <br> 2. Prudent: vegetables and fruit, soybean curd, fish and milk <br> Fatty: meat and fatty foods <br> Japanese: cooked rice for breakfast and miso soup <br> Salty: pickles, dried or salted fishes and salty foods <br> 3. $21 \%$ | Prudent: inverse association with breast cancer risk (Women in the highest quartile of the prudent dietary pattern scores, had a $27 \%$ decreased risk of breast cancer compared with those in the lowest ( $95 \% \mathrm{CI}: 0.63-0.84$, p for trend $<$ 0.0001 ). | * |
| (Kroenke et <br> al., 2005) | Cancer (Breast) | 1. 2 <br> 2. Prudent: fruits, vegetables, whole grains, legumes, poultry, and fish Western: refined grains, processed and red meats, desserts, high-fat dairy products, and French fries. | All patterns: Not Statistically Significant | * |
| (Velie et al., 2005) | Cancer (breast) | 1. 3 <br> 2. vegetable-fish/ poultry-fruit : vegetables and broiled or baked fish and chicken and low intakes of sweets and white bread beef/pork-starch: pork, beef, processed meat, French fries, and eggs and low intakes of bran cereal, skim milk, broiled or baked fish and chicken, and dark bread traditional southern : traditional rural southern US foods <br> 3. $12.5 \%$ | All patterns: Not Statistically Significant | vegetable-fish/poultry-fruit : was associated with higher education |
| $\begin{aligned} & \text { (Sieri et al., } \\ & 2004) \end{aligned}$ | Cancer (Breast) | 1. 4 <br> 2. Salad vegetables: raw vegetables and olive oil <br> Western: potatoes, red meat, eggs, butter, seed oil (as added fat) and cakes <br> Canteen: pasta, tomato sauce, and wine. Prudent: cooked vegetables, rice, poultry, fish, and low consumption of alcohol | salad vegetables : was associated with significantly lower breast cancer incidence $(\mathrm{RR}=0.66, \mathrm{CI} 95 \%=0.47,0.95$ )comparing highest with lowest tertile) with a significant linear trend $(\mathrm{P}=0.016)$ <br> Other patterns: Not Statistically Significant | * |
| $\begin{aligned} & \text { (Ronco et al., } \\ & 2006 \text { ) } \end{aligned}$ | Cancer (Breast) | 1. 6 <br> 2. Traditional: boiled meat, grains, cooked vegetables and tubers Healthy: white meat, raw vegetables, | Western: positively associated with breast risk cancer ( (OR 1.31, 95\% CI 1.13-1.51) <br> Traditional (OR 0.77, 95\% CI 0.64-0.93 | * |

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|  |  |  | cooked vegetables and total fruits <br> Western: fried meat, barbecue and <br> processed meat <br> Stew: boiled meat and legumes. <br> High-Fat: dairy foods and eggs and was labelled <br> Drinker: alcohol $58.3 \%$ | Healthy (OR 0.84, 95\% CI 0.73-0.98 and stew (OR 0.83, $95 \%$ CI $0.71-0.98$ diets were significantly protective. <br> Other patterns : Not Statistically Significant |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (De Stefani et <br> al., 2007) | Cancer ( Laryngeal carcinoma) |  | 6 <br> (Pattern 1)Traditional <br> (Pattern 2)Healthy: fish, fresh vegetables, total fruits and tea <br> Third pattern: cheese, butter, mayonnaise and custard <br> (Pattern 4)High - Fat: processed meat, soft drinks and negative loading with whole milk. <br> (Pattern 5)Drinker pattern : beer, wine and hard liquor and can be labelled as the drinker pattern <br> (Pattern 6)Western : fried meat, barbecued meat, processed meat and fried eggs 36.4\% | (Pattern 5) Drinker : directly associated with risk of laryngeal carcinoma (OR 3.8, 95\% CI 1.9-7.5, whereas the Pattern 2 <br> (Pattern 2) Healthy was protective to cancer (OR 0.6, $95 \%$ CI 0.3-1.2. <br> (Pattern 6) Western: significant increase in risk of laryngeal cancer 3.2 ( $95 \%$ CI 1.6-6.2.). <br> Other patterns: Not Statistically Significant | Pattern 1: positively associated with those who were frequently lived in a rural environment, and were less educated and smoked significantly more than those in the other patterns. <br> Pattern 2 was positively associated with education and high body mass and inversely associated with tobacco smoking. <br> Pattern 3 was directly associated with total energy and total fat intake. <br> Pattern 4 was inversely associated with tobacco smoking and b-carotene, <br> Pattern 5 was strongly associated with pack-years of tobacco smoking, black tobacco use and handrolled cigarettes preference. This pattern was also highly correlated with mate drinking. |
| (AgursCollins et al., 2009) | Cancer ( Breast) |  | 2 <br> Western: refined grains, high-fat dairy products, meat and processed meat, eggs, margarine, butter and mayonnaise, potato, French fries, sweets, soda, and snacks. Prudent: cruciferous and other vegetables, fruit, whole grains, cereals, beans, low-fat dairy products, fish, and poultry. $22 \%$ | Prudent: associated with lower breast cancer risk overall ( P for trend $=0.06$ ) highest vs. lowest quintile $(\mathrm{OR}=0.86$, 95\% CI: $0.68,1.08$ ). <br> Western: Not Statistically Significant | Western: positively associated with younger age, weight, lower education, smoking and drinking, less exercise <br> Prudent : positively associated with older age, higher educational status, less likelihood to smoke, higher levels of strenuous physical activity |
| $\begin{aligned} & \text { (Dixon et al., } \\ & 2004 \text { ) } \end{aligned}$ | Cancer (Colon) |  | 2. <br> Vegetarian: vegetables and legumes, citrus fruit and berries, pasta and rice, poultry and fish, and oil and salad dressings <br> PPP: pork, processed meats, potatoes, and coffee. <br> $5.6 \%$ in the men to $9.7 \%$ in the women | PPP: associated with an increased risk of colon cancer in women(quintile multivariate relative risk: $1.62 ; 95 \% \mathrm{CI}: 1.12$, 2.34; Pfor trend $<0.01$ ) <br> PPP: associated with an increased risk of colorectal cancer in men (quintile 4multivariate relative risk: $2.21 ; 95 \% \mathrm{CI}$ : 1.07, 4.57; $P$ for trend $<0.05$ ). <br> Other patterns: Not Statistically Significant | * |
| (Hughes et | Actinic Keratoses |  | 3 <br> the vegetable-fruit potato-sweets-meat | All the patterns: Not Statistically Significant | * |

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|  |  | 3. | fats, butter, French fries, high-fat dairy products, egg, organ meats, and sugar 14.5\% |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Robinson et <br> al., 2009) | * |  | 2 <br> Prudent :fruit, vegetables, oily fish and wholemeal cereals Traditional: vegetables, processed and red meat, fish and puddings | * | Prudent : positively associated with women, nonmanual classes and non-smoking <br> Traditional : positively associated with men, men and women who had partners and alcohol consumption |
| (Touvier et <br> al., 2009) | * |  | 3 <br> Processed/starchy meat: biscuits, bread, butter/cream, cakes, French fries, marmalade/honey, pulses, pizza/pies, potatoes, processed meat, rice, pasta, semolina, sandwiches, dairy-based sweet puddings, sugar/confectionary, low consumption of vegetables <br> Fruit/vegetables: breakfast cereals, eggs, fruits, ham, pulses, offal, olive oil, seafood, vegetable oil, vegetables, yoghurt/ cottage cheese <br> Alcohol/meat products: alcohol beverages, wine, meat, offal, poultry/rabbit, processed meat, appetizers, artificial sweeteners, coffee and French fries, as well as with low consumption of fruits, soup and marmalade/honey. |  | Fruit/vegetables : positively associated with supplement use, smoking, <br> Processed meat/Starchy foods and Alcohol/meat inversely associated with supplement use, smoking, |
| (Uusitalo et al., 2009) | * |  | 7 <br> Healthy: vegetables, fish , rice, eggs , meat dishes, creamy Fast food: processed meat, chocolate , sweets, sausage , eggs <br> Traditional bread: meat, cereals, bread Traditional meat : meat, sausage potatoes Low-fat: low fat cheese, meat and soft drinks <br> Coffee <br> Alcohol and butter: beer wine and butter | * | * |
| (Arkkola et al., 2008) | * |  | 6 <br> Healthy: Vegetables, pasta , eggs, fruits Fast food: processed and fried foods and snacks <br> Traditional bread/Traditional meat | * | * |

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\begin{tabular}{|c|c|c|c|c|c|}
\hline \& \& 3 \& Low-fat Coffee 29.5\% \& \& \\
\hline \begin{tabular}{l}
(Borland et \\
al., 2008)
\end{tabular} \& * \& \[
\begin{aligned}
\& 1 . \\
\& 2 .
\end{aligned}
\] \& \begin{tabular}{l}
2 \\
Prudent: vegetables, fruit, wholemeal bread, rice/pasta, yoghurt and breakfast High - energy: Puddings, cakes/biscuits, potatoes/chips, vegetables, fruit, red/processed meat, fish, eggs, oils and fullfat spreads
\end{tabular} \& * \& \begin{tabular}{l}
Prudent: associated with strenuous exercise taken and partnership status. \\
High-energy: associated with lower social class.
\end{tabular} \\
\hline \begin{tabular}{l}
(Crozier et \\
al., 2008)
\end{tabular} \& * \& \& \begin{tabular}{l}
2 \\
Prudent: fruit and vegetables, wholemeal bread, rice and pasta, yoghurt, cheese, fish and reduced-fat milk, low in white bread, added sugar, tinned vegetables, full-fat milkand crisps. \\
Western: processed meat, cakes and biscuits, puddings, Yorkshire puddings and savoury pancakes, chips, roast and boiled potatoes, sugar, sweets and chocolate.
\end{tabular} \& * \& * \\
\hline \begin{tabular}{l}
(Engeset et \\
al., 2009)
\end{tabular} \& * \& 1.
2.

3. \& | 6 |
| :--- |
| Traditional fish eaters Healthy |
| Average |
| Less fish |
| Less healthy |
| Western |
| Traditional bread eaters |
| Alcohol users |
| 23.7\% | \& * \& * <br>

\hline | (Keskitalo et |
| :--- |
| al., 2008) | \& * \& \[

$$
\begin{aligned}
& 1 . \\
& 2 .
\end{aligned}
$$

\] \& | 4 |
| :--- |
| Healthy: Vegetables, rice, chicken, |
| yogurt, fish |
| High-fat: Processed food, pizza , creamy |
| foods |
| Sweets: Salty snacks, chocolate |
| Meats: Sausage, meat | \& * \& <br>


\hline (Knudsen et al., 2008) \& * \& \& | 2 |
| :--- |
| Western : red and processed meat, high-fat dairy |
| Health Conscious: vegetables, fruits, poultry and fish. | \& * \& Western: positively associated with smoking Health Conscious: negatively associated with smoking <br>

\hline \[
$$
\begin{aligned}
& \text { (Lau et al., } \\
& \text { 2008) }
\end{aligned}
$$

\] \& * \& \& | 2 |
| :--- |
| Traditional: high loadings on pate' or highfat meat for sandwiches, mayonnaise salads, | \& * \& * <br>

\hline
\end{tabular}

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|  |  | red meat, potatoes, butter and lard low-fat fish, low-fat meat for sandwiches, and sauces. <br> Modern: high loadings on vegetables, fruit, mixed vegetable dishes, vegetable oil and vinegar dressing, poultry, and pasta, rice and wheat kernels. |  |  |
| :---: | :---: | :---: | :---: | :---: |
| (Northstone \& Emmett, 2008a) | * | 1. 2 patterns consistently over time <br> 2. Processed <br> Traditional <br> Health conscious | * | * |
| (Northstone, <br>  <br> Rogers, <br> 2008a) | * | 1. 5 <br> 2. Health Conscious: salad, fresh fruit, rice, pasta, fish, pulses, and non-white bread Traditional: vegetables and to some extent red meat and poultry <br> Processed: meat pies, sausages and burgers, fried foods, pizza and chips <br> Confectionery: chocolate, sweets, biscuits, cakes and other pudding <br> Vegetarian: meat substitutes, pulses, nuts and herbal tea <br> 3. $31 \cdot 3 \%$ | * | Health conscious: positively associated with increasing education and age and non-white women and negatively associated with increased parity, single, non-working women, those who smoked and who were overweight pre-pregnancy. <br> Processed: Opposite associations |
| (Northstone, <br>  <br> Rogers, <br> 2008b) | * | Same as (Northstone, Emmett \& Rogers, 2008a) | Same as (Northstone, Emmett \& Rogers, 2008a) | Same as (Northstone, Emmett \& Rogers, 2008a) |
| (Northstone \& Emmett, 2008b) | * | Same as (Northstone, Emmett \& Rogers, 2008a) | Same as (Northstone, Emmett \& Rogers, 2008a) | Same as (Northstone, Emmett \& Rogers, 2008a) |

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$\left.\begin{array}{|l|l|ll|l|l|}\hline & & & \begin{array}{l}\text { potatoes, meats, savoury pies and } \\ \text { cruciferous vegetables } \\ \text { Dieting: low-fat dairy products, low-sugar } \\ \text { soda; low intake of butter and sweet baked } \\ \text { products. }\end{array} \\ \text { Low meat: baked beans, pizza and soy } \\ \text { foods; low intakes of meat, other fish and } \\ \text { seafood, and poultry }\end{array}\right]$

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| Vohl, 2006) |  | 3 | $\begin{aligned} & \hline \text { other seafood } \\ & 19.9 \% \text { and } 17.2 \% \end{aligned}$ |  | Prudent : positively associated with physical activity and negatively associated with BMI and |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Schulze et <br> al., 2006) | * |  | 2 <br> Prudent: fruits, vegetables Western: red and processed meats, refined grains, sweets and desserts, and potatoes whole grains, fish, poultry, and salad dressing | * | * |
| (Weismayer, Anderson \& Wolk, 2006) | * |  | 3 <br> Healthily: fruits, tomatoes, vegetables, cereal, and fish Western : meat, processed meat, fried potatoes, soft drinks, and sweets Alcohol: beer, wine, and liquor consumption as well as snacks | * | * |
| (Engeset et <br> al., 2005) | * | 3 | 6 <br> Traditional fish eaters Healthy Average <br> Less fish <br> Less healthy <br> Western Traditional bread eaters Alcohol users 23.7\% | * | Traditional fish eaters and the Traditional bread eaters: positively associated with lower income and lower education. <br> Healthy and the alcohol users : positively associated with higher income and south and east location <br> Alcohol users: positively associated with current smokers. |
| $\begin{aligned} & \text { (Fung et al., } \\ & 2005 \text { ) } \end{aligned}$ | * |  | 2 <br> Prudent: fruits, vegetables, whole grains, fish and poultry Western: processed and red meats, refined grains, sweets and desserts. | * | * |
| (Hoffmann et <br> al., 2005) | * |  | 2 <br> First PCA: potatoes, vegetables, legumes, bread, all types of meat, eggs, sauces and soups <br> Second PCA: vegetables, fruits, dairy products, other cereals, vegetable oils and non-alcoholic beverages, low consumption of alcoholic beverages other than wine. $11 \cdot 3$ and $8.4 \%$ | * | * |
| (Marchioni et <br> al., 2005) | * |  | 2 <br> Prudent: fruits and vegetables <br> Traditional: cereals and snacks <br> Snacks: dairy products and processed meat. | * | * |

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| (Mikkila et al., 2005) | * |  | 3 <br> Traditional Finnish foods: rye, potatoes, milk, butter, sausages and coffee, and low consumption with fruit, berries and dairy products other than milk. <br> Health-conscious foods: vegetables, legumes and nuts, tea, rye, cheese and other dairy products, and also alcoholic beverages. | * | * |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Northstone \& Emmett, 2005) | * |  | 3. <br> Junk : high-fat processed foods (sausages, burgers, coated poultry) and snack foods high in fat and/or sugar (such as crisps, sweets, chocolate, ice lollies and ice creams) <br> Traditional: meat and vegetables and Vegetarian: meat substitutes, pulses and nuts. | * | Junk: was significantly more likely in white children, where maternal education level was low and where the child had more siblings. <br> Traditional: was more likely in girls, where the mother had a partner and in no vegetarians (both mother and child). <br> Vegetarian: was more likely with increasing levels of education and increasing maternal, age. |
| $\begin{aligned} & \text { (Park et al., } \\ & 2005 \text { ) } \end{aligned}$ | * |  | 3 <br> Fat and Meat: discretionary fat, meat, eggs, and cheese. <br> Vegetables, Fruit and Milk: milk and yogurt and fruit groups. 63\% | * | Fat and Meat: positively associated with current smokers <br> Vegetables and Fruit and Milk : positively associated with physical activity |
| $\begin{aligned} & \text { (Perrin et al., } \\ & 2005 \text { ) } \end{aligned}$ | * |  | 2 <br> Western: sugar and sweets, grains, butter, added fats, eggs, dairy products, potatoes, cheese and fruit. <br> Prudent: fruit, vegetables, olive oil, dairy products and fish $26 \cdot 7 \%$ | * | Prudent: positively associated with region, educational and income tax levels, leisure-time physical activity and smoking status. |
| (Uusitalo et al., 2005) | * |  | 3 <br> Western: <br> Bread/butter Traditional High protein and Margarine/milk | * | Western : positively related with people who were younger, educated and wealthier subjects |
| (Yang, <br>  <br> Song, 2005) | * |  | 3 <br> Vegetable/fruit <br> Traditional Korean <br> Acculturated American | * | Traditional Korean: was negatively associated with length of residence in the U.S. for both men and women ( $p<0.01$ ). |
| $\begin{aligned} & \text { (Corrao et al., } \\ & 2004 \text { ) } \end{aligned}$ | * |  | 2 PC1 PC2 $75 \%$ | * | * |

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*no information provided

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Table D. Foods items which correlated >0.3 or <-0.3 with a "western" or a "Prudent" dietary pattern in different studies (Food items that didn't correlated $>0.3$ or <- 0.3 with any of the two patterns excluded from the table; Sorted alphabetically with year of publication and author; Studies that didn't identify a "Prudent" or "western" pattern were excluded from the table)

|  | Western <br> (foods that deemed to constitute the pattern with corresponding principal component loadings) | Prudent <br> (foods that deemed to constitute the dietary pattern with corresponding principal component loadings) |
| :---: | :---: | :---: |
| (Agurs-Collins et al., 2009) | ```eggs 0.41 french fries 0.55 high-fat dairy products 0.46 margarine, butter, and mayonnaise 0.40 potato 0.36 processed meat 0.62 refined grains 0.47 snacks 0.45 soda 0.42 sweets 0.47 total meat 0.65``` | beans 0.49 <br> cruciferous vegetables 0.65 <br> fish 0.48 <br> fruits 0.61 <br> juice 0.34 <br> low-fat dairy products 0.39 <br> other vegetables 0.75 <br> pasta 0.35 <br> poultry 0.36 <br> soup 0.41 <br> tomatoes 0.48 <br> whole grains 0.54 |
| (Ambrosini et al., 2009) | cakes, biscuits 0.34 <br> confectionery 0.46 <br> crisps 0.39 <br> full fat dairy products 0.30 <br> potato, fried e.g. french fries 0.39 <br> potato, not fried 0.34 <br> processed meats 0.41 <br> red meat 0.46 <br> refined grains 0.42 <br> sauces and dressings 0.34 <br> soft drinks 0.37 <br> takeaway foods 0.53 |  |
| (Bakolis et al., 2010) | butter beans/broad beans <br> roast potatoes 0.30 <br> ham 0.31 <br> ice cream 0.31 | avocado 0.37  <br> bean sprouts 0.30 244 <br> broccoli 0.38  <br> carrots 0.33  |

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|  | ```pork - roast, chops 0.32 pork stew, casserole 0.32 omelette/scrambled egg 0.32 fruit pies, tarts, crumbles 0.32 beef stew, casserole, mince, curry 0.34 sponge cakes 0.34 fried fish in batter/breadcrumb 0.35 baked beans 0.36 chocolate biscuits 0.36 sandwich/cream biscuits 0.36 corned beef, spam, luncheon meat 0.37 white bread and rolls 0.39 bacon 0.40 fried egg 0.40 milk chocolate 0.40 bread crumbed e.g. chicken nuggets 0.41 crisps 0.41 chocolate snack bars 0.44 meat pizza 0.46 other fried snacks 0.46 chips 0.47 sausages - beef, pork 0.48 beef burger, hamburger 0.53 pies/pasties/sausage rolls/meat 0.53``` | celery 0.30 <br> cheese - cheddar, brie, edam 0.31 <br> courgettes, marrow, squash 0.47 <br> couscous 0.36 <br> currants, raisins, sultanas 0.40 <br> french type dressing 0.54 <br> fresh oily fish 0.40 <br> garlic 0.49 <br> green beans, runner beans 0.37 <br> leeks 0.42 <br> lettuce 0.46 <br> mayonnaise 0.35 <br> mixed bean casserole/ratatouille 0.45 <br> mushrooms 0.44 <br> onions 0.49 <br> parsley - flat leaf 0.37 <br> parsnips and turnips 0.34 <br> peppers - red, green, yellow 0.52 <br> spinach 0.38 <br> sweetcorn 0.39 <br> tinned oily fish 0.33 <br> tofu - bean curd 0.30 <br> tomatoes - raw, canned, sauce 0.39 <br> vegeburgers 0.31 <br> vegetable - lasagne/moussaka 0.43 <br> vegetable pies/samosas 0.39 <br> vegetable pizza 0.46 <br> watercress, mustard and cress <br> white fish not fried 0.38 <br> white pasta 0.42 <br> wholemeal bread and rolls 0.33 <br> wholemeal pasta 0.33 |
| :---: | :---: | :---: |
| (Deshmukh-Taskar et al., 2009) | condiments 0.40 <br> dishes with cheese 0.58 <br> eggs 0.39 <br> french fries 0.53 | $\begin{array}{\|l\|} \hline 100 \% \text { fruit juices } 0.43 \\ \text { clear soups } 0.36 \\ \text { cruciferous vegetables } 0.70 \\ \text { dark-yellow vegetables } 0.70 \\ \hline \end{array}$ |

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|  | high-fat dairy products 0.53 processed meats 0.59 <br> red meats 0.50 <br> refined grains 0.43 <br> snacks 0.53 <br> sweetened beverages 0.44 <br> sweets and desserts 0.54 | fruits 0.64 <br> green leafy vegetables 0.69 <br> legumes 0.61 <br> low-fat dairy products 0.36 <br> low-fat salad dressings 0.49 <br> other vegetables -0.74 <br> poultry 0.40 <br> tomatoes 0.58 <br> whole grains -0.46 |
| :---: | :---: | :---: |
| (Imamura et al., 2009) | high-fat dairy 0.32 <br> high-fat dairy desserts 0.35 <br> nuts and seeds 0.30 <br> eggs 0.39 <br> processed meat 0.61 <br> meat 0.55 <br> refined grains 0.34 <br> chocolate 0.38 <br> sweet baked goods 0.460 .37 <br> chowder/cream soup 0.32 <br> soda 0.38 <br> pizza, sandwich, casserole 0.36 <br> potato or corn chips 0.48 <br> fried foods 0.48 | reduced-fat dairy 0.39 <br> fruits 0.50 <br> whole grains 0.35 <br> refined-grain cereal 0.33 <br> sweet baked goods 0.37 <br> miscellaneous sweets 0.42 |
| (Kesse-Guyot et al., 2009) |  | fruits 0.50 <br> tea 0.34 <br> butter 0.47 <br> vegetable oil 0.30 <br> breakfast cereals 0.39 <br> vegetables 0.49 <br> fish and seafood 0.30 <br> reduced-fat products 0.44 |
| (Lutsey et al., 2009) | processed meat, noncereal whole grains added fats oils | vegetables <br> fruit <br> poultry |

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| (Oddy et al., 2009) | takeaway foods 0.53 confectionery 0.46 red meat 0.46 refined grains 0.42 processed meats 0.41 potato, fried e.g. french fries 0.39 crisps 0.39 soft drinks 0.37 cakes, biscuits 0.34 potato, not fried 0.34 sauces and dressings 0.34 full fat dairy products 0.30 | yellow or red vegetables 0.56 <br> leafy green vegetables 0.49 <br> tomato 0.49 <br> cruciferous vegetables 0.48 <br> other vegetables 0.66 <br> fresh fruit 0.48 <br> legumes 0.43 <br> wholegrain 0.39 <br> fish, steamed, grilled or tinned 0.33 |
| :---: | :---: | :---: |
| (Paradis et al., 2009) | refined grains 0.68 <br> french fries 0.61 <br> red meats 0.57 <br> condiments 0.50 <br> processed meats 0.50 <br> regular soft drinks 0.48 <br> pizza 0.44 <br> snacks 0.37 <br> potatoes other than french fried 0.35 <br> legumes -0.31 <br> fruits -0.44 | non-hydrogenated fat 0.56 <br> vegetables 0.52 <br> eggs 0.46 <br> fish and other seafood 0.45 <br> wine 0.44 <br> coffee 0.42 <br> regular dairy products 0.37 desserts <br> whole grains 0.32 |
| (Robinson et al., 2009) | fruit <br> vegetables <br> oily fish <br> wholemeal cereals |  |
| e3n study |  |  |
| (Varraso et al., 2009) (Touvier et al., 2009) | condiments and sauces 0.32 onions, garlic 0.73 <br> dough and pastry 0.70 <br> cream desserts 0.62 <br> ice cream 0.60 <br> processed meats 0.550 .33 <br> cakes, pies and pastries 0.45 <br> pasta, rice and grain 0.40 | fruity vegetables 0.89 <br> root vegetables 0.85 <br> cabbages 0.79 <br> mushrooms 0.73 <br> grain and peas 0.72 <br> leafy vegetables (except <br> cabbages) 0.71 <br> stalk vegetables 0.70 |

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$\left.\begin{array}{|l|l|l|}\hline & \begin{array}{l}\text { potatoes and other tubers } 0.31 \\ \text { egg } 0.30\end{array} & \begin{array}{l}\text { fruits with beta carotene } 0.61 \\ \text { fruits with citric } 0.60 \\ \text { condiments and sauces } 0.42 \\ \text { red meat } 0.34 \\ \text { poultry } 0.33 \\ \text { blue fish } 0.32\end{array} \\ \hline \begin{array}{l}\text { (Kesse, Clavel-Chapelon \& Boutron- } \\ \text { Ruault, 2006) }\end{array} & \begin{array}{l}\text { potatoes } 0.45 \\ \text { pizza and pies } 0.48 \\ \text { sandwiches } 0.32 \\ \text { legumes } 0.32 \\ \text { sweets } 0.42 \\ \text { cakes } 0.41 \\ \text { pasta } 0.63\end{array} \\ \text { rice } 0.55 \\ \text { bread } 0.37 \\ \text { processed meat } 0.39\end{array}\right)$.

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|  |  | ```white bread 0.45 eggs 0.38 potato crisps 0.36 potato chips (french fries) 0.31 meat pie 0.46 hamburger 0.40 beef 0.42 lamb 0.32 pork 0.37 bacon 0.37 sausages 0.48 full cream milk 0.31 beer (full alcohol) 0.30 fish, fried or takeaway 0.41``` |
| :---: | :---: | :---: |
| southampton women's survey study (sws) |  |  |
| (Borland et al., 2008) |  | vegetables <br> fruit <br> wholemeal bread <br> rice/pasta <br> yoghurt and breakfast cereals <br> lower intakes of <br> white bread <br> roast potatoes/ chips <br> red/processed meat <br> full-fat milk <br> full-fat spread, crisps <br> confectionery <br> sugar <br> tea/coffee <br> yorkshire puddings/ pancake <br> tinned vegetables <br> cakes and biscuits <br> soft drinks <br> high energy |
| (Crozier et al., 2006) |  | rice and pasta 0.21 white bread -0.22 |

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|  |  | wholemeal bread 0.23 <br> full-fat milk -0.20 <br> processed meat -0.21 <br> salad vegetables 0.25 <br> other vegetables 0.23 <br> vegetable dishes 0.21 <br> chips and roast potatoes -0.27 <br> other fruit 0.23 <br> sugar -0.23 |
| :---: | :---: | :---: |
| (Campbell, Sloan \& Kreiger, 2008) <br> Different Gender <br> Men Women | ```potatoes: baked, broiled, mashed 33* french fries or fried potatoes 34* 34* white bread 35* 36* beef, pork, or lamb as a main dish 25* 36* hamburger 40* 46* hot dogs 37* 40* hamburger 40* 46* hot dogs \(37 * 40^{*}\) luncheon meats (salami, bologna) 30* 30* smoked meat or corned beef \(26^{*} 25^{*}\) bacon 34* 45* sausage 31*45* eggs 33* 41* cheese 29* cake 42* doughnuts, pastry 46* 36* pies 39* 35* ice cream 35* \(28^{*}\) chocolate 28* 29* potato chips 32* \(28^{*}\) butter on bread or vegetables 26*29* mayonnaise or salad dressing on bread or in salads 27* 29*``` | ```tomato or vegetable juice \(26^{*}\) apples or pears 43* 38* oranges 34* 38* bananas 33* 40* cantaloupe 29* \(35^{*}\) other fruit, fresh or canned 40*39* tomatoes 41* 41* carrots 40* 54* broccoli 52* \(58^{*}\) cabbage, cauliflower, brussels sprouts 47* 51* spinach or other greens 44*51* yellow squash \(26 * 44 *\) other vegetables \(35^{*} 47^{*}\) soups with vegetables 32 * \(30^{*}\) sweet potatoes \(27^{*}\) baked beans or lentils 30* rice 27* chicken or turkey 32* fish: fresh, frozen canned 33* 31*``` |
| (De Stefani et al., 2008a) | red meat 0.49 <br> poultry -0.55 <br> fish -0.52 <br> wine 0.41 <br> cheese - 0.34 | desserts 0.45 <br> french bread -0.40 <br> raw vegetables 0.41 <br> cooked vegetables 0.55 <br> citrus fruits 0.45 |

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|  | fried eggs 0.34 <br> wine 0.41 | other fruits 0.51 |
| :---: | :---: | :---: |
| nurses health study i and ii (nhs) |  |  |
| (Heidemann et al., 2008) Different time points 19841986199019941998 | ```refined grains 0.580 .520 .460 .50 processed meat 0.570 .580 .600 .580 .58 red meat 0.550 .570 .600 .610 .62 french fries 0.470 .480 .470 .470 .48 condiments 0.450 .320 .350 .320 .36 sweets and desserts0.43 0.490 .430 .410 .36 potatoes 0.390 .360 .300 .290 .33 high-fat dairy 0.370 .420 .480 .470 .45 pizza 0.350 .350 .360 .340 .39 mayonnaise 0.340 .350 .350 .330 .33 high-sugar beverages 0.320 .320 .320 .290 .30 eggs 0.300 .330 .420 .440 .40 margarine 0.290 .270 .280 .310 .25 snacks 0.280 .320 .280 .320 .32 butter 0.270 .290 .330 .330 .31 soups 0.220 .290 .310 .340 .32``` | other vegetables 0.680.72 0.690 .710 .71 <br> green, leafy vegetables 0.650 .640 .640 .640 .63 <br> cruciferous vegetables 0.610 .600 .61 0.630.62 <br> legumes 0.590 .560 .580 .590 .57 <br> dark-yellow vegetables 0.580 .630 .650 .640 .62 <br> fruit 0.580 .590 .570 .580 .59 <br> fish 0.510 .530 .510 .470 .50 <br> tomatoes 0.460.550.49 0.520.53 <br> poultry 0.440 .410 .420 .340 .400 .17 <br> whole grains 0.390 .390 .410 .400 .40 - <br> salad dressing0.36 0.380 .340 .310 .33 <br> low-fat dairy $0.32 \quad 0.320 .320 .33$ <br> olive oil na na 0.310 .320 .39 |
| (Varraso et al., 2007a) | refined grains 0.74 <br> desserts and sweets 0.60 <br> cured meats 0.52 <br> red meats 0.52 <br> french fries 0.44 <br> condiments 0.40 <br> potatoes 0.39 <br> pizza 0.36 <br> full-fat dairy products 0.35 <br> sweetened beverages 0.32 <br> mayonnaise 0.31 <br> margarine 0.30 | other vegetables 0.68 <br> leafy vegetables 0.63 <br> cruciferous vegetables 0.61 <br> fruit 0.60 <br> yellow vegetables 0.60 <br> legumes 0.55 <br> fish 0.50 <br> tomatoes 0.45 <br> poultry 0.43 <br> whole-grain products 0.41 <br> low-fat dairy products 0.35 <br> garlic 0.35 <br> salad dressing 0.33 |
| (Schulze et al., 2006) Different time points 199119951999 | red meats $0.610 .55 \quad 0.62$ <br> processed meats 0.580 .540 .49 <br> french fries 0.500 .510 .54 <br> refined grains 0.470 .430 .51 | other vegetables 0.680 .700 .69 green, leafy vegetables 0.670 .650 .61 dark-yellow vegetables 0.630 .590 .55 fruit 0.620 .610 .59 |

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|  | sweets and desserts 0.420 .430 .39 <br> potatoes 0.410 .370 .43 <br> eggs 0.390 .390 .20 <br> snacks 0.390 .330 .33 <br> high-fat dairy products 0.360 .340 .31 <br> margarine 0.360 .380 .32 <br> pizza 0.340.39 0.40 <br> mayonnaise0.34 0.360 .31 <br> sugar-sweetened soft drinks 0.300 .280 .30 | cruciferous vegetables 0.580 .580 .57 <br> tomatoes 0.540 .440 .46 <br> legumes 0.500 .530 .50 <br> fish and other seafood0.44 0.390 .43 <br> oil and vinegar salad dressing0.43 0.500 .55 <br> whole grains 0.430 .430 .41 <br> poultry 0.400 .240 .31 <br> garlic 0.380 .410 .43 <br> water 0.380 .390 .39 <br> condiments 0.310 .30 - |
| :---: | :---: | :---: |
| (Fung et al., 2005) | refined grains 0.74 <br> desserts and sweets 0.60 <br> processed meats 0.52 <br> red meats 0.52 <br> french fries 0.44 <br> condiments 0.40 <br> potatoes 0.39 <br> pizza 0.36 <br> full-fat dairy products 0.35 <br> sweetened beverages 0.32 <br> mayonnaise 0.31 <br> margarine 0.30 | other vegetables 0.68 <br> leafy vegetables 0.63 <br> cruciferous vegetables 0.61 <br> fruit 0.60 <br> yellow vegetables 0.60 <br> legumes 0.55 <br> fish 0.50 <br> tomatoes 0.45 <br> poultry 0.43 <br> whole grain products 0.41 <br> low-fat dairy products 0.35 <br> salad dressings 0.33 <br> garlic 0.35 |
| (Kroenke et al., 2005) | refined grains processed and red meats desserts high-fat dairy products french fries | fruits <br> vegetables <br> whole grains <br> legumes <br> poultry <br> fish |
| (Fung et al., 2004a) Different time points 1984198619901994 | refined grains 0.580 .570 .520 .44 processed meats 0.560 .570 .580 .57 red meats 0.560 .560 .600 .61 french fries 0.470 .470 .460 .47 condiments 0.440 .330 .360 .29 desserts and sweets 0.430 .490 .460 .46 potatoes 0.410 .390 .340 .34 | other vegetables 0.690 .750 .680 .68 leafy vegetables 0.680 .680 .670 .60 cruciferous vegetables 0.610 .590 .610 .63 yellow vegetables 0.560 .600 .640 .66 fruits 0.550 .550 .550 .62 fish 0.520 .540 .520 .43 legumes 0.510 .490 .550 .60 |

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|  | ```full fat dairy products 0.360 .400 .450 .43 pizza 0.350 .340 .350 .33 sweetened beverages _0.33 0.340 .330 .33 margarine 0.330 .310 .330 .34 mayonnaise 0.320 .330 .340 .27 eggs 0.290 .320 .410 .41 snacks 0.280 .310 .290 .33 butter 0.240 .250 .290 .27 cream soups- 0.300 .320 .35 water n/an/a 0.300 .35 olive oil n/a n/a 0.350 .21``` | tomatoes 0.480 .570 .510 .46 poultry 0.450 .420 .420 .32 garlic $0.41 \mathrm{n} / \mathrm{a} \mathrm{n} / \mathrm{a} 0.26$ salad dressings 0.390 .410 .380 .24 whole grains 0.350 .340 .350 .42 low-fat dairy products 0.280 .260 .280 .37 fruit juice 0.210 .220 .190 .26 nuts 0.160 .200 .15 |
| :---: | :---: | :---: |
| (Lopez-Garcia et al., 2004) Different time points 19861990 1986/1990 | processed meats 0.570 .580 .59 <br> refined grains 0.570 .520 .58 <br> red meats 0.560 .600 .61 <br> sweets and desserts 0.490 .450 .49 <br> french fries 0.480 .460 .49 <br> high-fat dairy products 0.400 .460 .45 <br> potatoes 0.380 .330 .39 <br> pizza 0.340 .360 .37 <br> sugar-containing beverages 0.340 .330 .36 <br> mayonnaise and other creamy salad dressings <br> 0.330 .340 .36 <br> condiments 0.330 .370 .38 <br> eggs 0.320 .410 .38 <br> snacks 0.310 .290 .31 <br> margarine 0.300 .320 .34 <br> cream soup 0.300 .310 .32 | other vegetables 0.730 .680 .73 <br> green, leafy vegetables 0.660 .650 .68 <br> dark-yellow vegetables 0.620 .650 .66 <br> cruciferous vegetables 0.590 .610 .62 <br> fruit 0.580 .580 .60 <br> tomatoes 0.560 .490 .55 <br> fish and other seafood 0.530 .510 .55 <br> legumes 0.510 .570 .55 <br> poultry 0.420 .420 .44 <br> salad dressings 0.400 .340 .39 <br> whole grains 0.380 .410 .41 <br> low-fat dairy products 0.300 .320 .32 |
| (Iqbal et al., 2008) | dairy 0.56 <br> nuts 0.29 <br> glv 0.32 <br> raw vegetables other than glv 0.63 <br> fruits 0.68 <br> desserts 0.40 | eggs 0.44 <br> meats 0.39 <br> fried foods 0.63 <br> salty foods 0.61 <br> sugar 0.32 |
| (Kubo et al., 2008) | beef 59 french fries 53 hamburger 52 |  |

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\(\left.$$
\begin{array}{|l|l|l|}\hline & \begin{array}{l}\text { pizza 49 } \\
\text { refried beans 49 } \\
\text { mustard 47 } \\
\text { tacos 45 } \\
\text { white potatoes 44 } \\
\text { pork chop 44 } \\
\text { pork spareribs 44 } \\
\text { chili 43 } \\
\text { hot dog 43 } \\
\text { salty snacks 43 } \\
\text { spaghetti 40 } \\
\text { fried chicken 38 } \\
\text { bacon 36 } \\
\text { sausage 36 } \\
\text { jelly 35 }\end{array} & \\
\hline \text { (Lutsey, Steffen \& Stevens, 2008) } & \begin{array}{l}\text { refined-grain bread, cereal, rice, and pasta } 0.63 \\
\text { processed meat } 0.63 \\
\text { fried foods } 0.61 \\
\text { red meat } 0.57 \\
\text { eggs } 0.48 \\
\text { refined-grain desserts } 0.43 \\
\text { soda and sweetened beverages } 0.41 \\
\text { cheese and whole milk } 0.38 \\
\text { legumes } 0.35 \\
\text { sweets/candy } 0.30 \\
\text { fat } 0.30\end{array} & \\
\hline \text { (Murtaugh et al., 2008) } & \begin{array}{l}\text { low-fat dairy } \\
\text { whole grains } \\
\text { fruit and fruit juice } \\
\text { legumes } \\
\text { vegetables } \\
\text { soups }\end{array} & \begin{array}{l}\text { cruciferous vegetables } 0.62 \\
\text { carotenoid vegetables } 0.60 \\
\text { fruit (no juice) } 0.58 \\
\text { other vegetables } 0.52 \\
\text { fish and seafood } 0.46 \\
\text { poultry } 0.43 \\
\text { dark leafy vegetables } 0.43\end{array}
$$ <br>
whole grains 0.40 <br>

tomatoes 0.39\end{array}\right]\)| legumes 0.34 |
| :--- |
| low-fat dairy 0.31 |

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|  |  | high-fat and high-sugar desserts |
| :---: | :---: | :---: |
| 54 universities in japan |  |  |
| (Okubo et al., 2008) | breads-0.28 <br> confectionary -0.35 <br> meats 0.60 <br> fats and oils 0.58 <br> seasonings 0.50 <br> processed meats 0.45 <br> eggs 0.30 |  |
| (Okubo et al., 2007) | confectionaries -0.33 <br> fats and oils 0.60 <br> meats 0.58 <br> seasonings 0.51 <br> processed meats 0.46 <br> eggs 0.33 |  |
| (Romaguera et al., 2008) | beef (0.51) <br> lamb (-0.61) <br> common bread ( $0 \cdot 42$ ) <br> bollo and tortilla $(-0.52)$ <br> chicken ( 0.42 ) <br> animal fat $(-0 \cdot 46)$ <br> fruit (0.38) <br> creole potatoes $(-0 \cdot 43)$ <br> sugary drinks ( $0 \cdot 29$ ) <br> mote $(-0 \cdot 30)$ <br> common potatoes $(0 \cdot 25)$ <br> herbal teas $(-0 \cdot 32)$ <br> yoghurt ( $0 \cdot 25$ ) <br> llama (-0.27) <br> green beans ( $0 \cdot 23$ ) <br> vegetables ( $-0 \cdot 20$ ) <br> sweet and milky desserts $(0 \cdot 21)$ |  |
| (Sadakane et al., 2008) |  | high -fat products -0.32 <br> bread 0.59 <br> butter 0.53 <br> rice -0.51 |

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|  |  | $\begin{aligned} & \text { salty products }-0.48 \\ & \text { miso-soup }-0.43 \\ & \text { yoghurt } 0.42 \\ & \text { fruits } 0.30 \\ & \text { milk } 0.38 \end{aligned}$ |
| :---: | :---: | :---: |
| (De Stefani et al., 2007) | fried meat 0.62 <br> barbeque meat 0.56 <br> poultry -0.31 <br> processed meat 0.32 <br> boiled eggs 0.32 <br> fried eggs 0.39 |  |
| (Hirose et al., 2007) |  | carrot 66 <br> green leafy vegetables 65 <br> potato 53 <br> pumpkin 52 <br> cabbage 48 <br> soy bean curd (tofu) 46 <br> fruit 45 <br> raw vegetables 45 <br> cooked/raw fish 30 <br> milk 30 <br> lettuce 36 |
| (Kim et al., 2007) | flour and bread pizza and hamburgers snacks and cereals sugars and sweets meats beverages |  |
| (Meyerhardt et al., 2007) | high-fat dairy 0.67 low-fat dairy 0.64 refined grains 0.60 condiments 0.51 red meat 0.53 sweets and desserts 0.53 margarine 0.50 processed meat 0.45 | vegetables 0.72 <br> leafy vegetables 0.71 <br> yellow vegetables 0.67 <br> cruciferous vegetables 0.65 <br> legumes 0.56 <br> fruit 0.55 <br> light salad dressing 0.48 <br> tomatoes 0.460 .36 |

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|  | ```potatoes 0.170 .45 regular mayonnaise 0.35 butter 0.33 french fries -0.16 0.37 eggs 0.30 snacks 0.36 nuts 0.30``` | garlic 0.39 <br> fish 0.46 <br> poultry 0.37 <br> fruit juice 0.35 <br> whole grains 0.32 <br> low-fat mayonnaise 0.31 |
| :---: | :---: | :---: |
| (Murtaugh et al., 2007) | high-cholesterol eggs 0.32 <br> high-fat dairy 0.41 <br> refined grains (regular) 0.60 <br> refined-grain snacks (regular) 0.43 <br> refined-grain cereals (regular) 0.33 <br> gravy and sauces 0.54 <br> sauces, tomato-based 0.39 <br> fast-food vegetables (french fries) 0.58 <br> fast-food beef sandwiches, hamburgers 0.59 <br> fast-food chicken 0.50 <br> bacon, sausage, cold cuts 0.54 <br> potatoes 0.55 <br> margarine 0.40 <br> polyunsaturated oils 0.39 <br> sugar 0.43 <br> high-fat, high-sugar desserts 0.41 <br> no-fat, high-sugar desserts 0.30 <br> meats 0.54 <br> mexican meats 0.30 | low-fat dairy 0.35 <br> whole grains (regular) 0.34 <br> whole-grain cereals 0.44 <br> orange, grapefruit, citrus juices 0.40 <br> fruit juices other than citrus 0.50 <br> canned fruit 0.46 <br> dried fruit 0.51 <br> soups (broth or cream based) 0.37 <br> soy beans, tofu 0.32 <br> salad greens, lettuce 0.30 <br> legumes, beans 0.36 <br> nuts 0.37 <br> tea, herbal 0.32 <br> fresh fruit 0.60 <br> vegetables 0.47 |
| (Sant et al. 2007) | ```potatoes ravioli red and processed meat eggs butter seed oil cakes.``` |  |
| health professionals study (hps) |  |  |
| (Qi et al., 2009) | processed meat, red meat, butte | vegetables fruit |

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|  | high-fat dairy products eggs refined grains | legumes whole grains fish poultry |
| :---: | :---: | :---: |
| (Wu et al., 2004a) (Varraso et al., 2007b) | red meat $0.64 \quad 0.670 .67$ <br> processed meat 0.600 .620 .61 <br> refined grains 0.460 .390 .34 <br> french fries 0.460 .500 .50 <br> high-fat dairy $0.44 \quad 0.50 \quad 0.49$ <br> sweets and desserts $0.41 \quad 0.42 \quad 0.41$ <br> eggs 0.410 .480 .46 <br> condiments $0.36-0.37$ <br> high-sugar drinks $0.33 \quad 0.32 \quad 0.29$ <br> snacks 0.330 .310 .30 <br> mayonnaise 0.340 .340 .35 <br> butter 0.320 .350 .36 | other vegetables $0.74 \quad 0.700 .69$ <br> dark-yellow vegetables 0.610 .630 .63 <br> cruciferous vegetables $0.59 \quad 0.62 \quad 0.62$ <br> green, leafy vegetables $0.640 .60 \quad 0.61$ <br> legumes $0.58 \quad 0.58 \quad 0.59$ <br> fruit 0.560 .560 .55 <br> tomatoes $0.55 \quad 0.50 \quad 0.51$ <br> fish 0.470 .450 .42 <br> whole grains $0.360 .36 \quad 0.36$ <br> poultry 0.340 .340 .28 <br> other salad dressing $0.35 \quad 0.36 \quad 0.24$ |
| (Michaud et al., 2005) <br> Different Studies <br> HPFS NHS | red meat . 64.53 <br> processed meat 60.54 <br> french fries . 45.45 <br> refined grains .45 .72 <br> high-fat dairy products. 44.36 <br> condiments . 42.44 <br> eggs .42 .26 <br> sweets and desserts . 38 . 58 <br> mayonnaise . 34.33 <br> snacks . 33.30 <br> sugar drinks . 32 . 31 <br> butter . 32.25 <br> margarine 28.28 <br> potatoes . 28.38 | vegetables . 73.69 <br> leafy vegetables . 63.66 <br> yellow vegetables . 61.58 <br> cruciferous vegetables . 59 . 61 <br> legumes . 59 . 53 <br> fruit 58.58 <br> tomatoes . 54.46 <br> fish .47 .51 <br> whole grains . 36.39 <br> poultry. 34.44 <br> salad dressing 33.37 <br> low-fat dairy products - . 32 |
| (Pala et al., 2006) |  | other vegetables 0.69 <br> legumes (pulses) 0.61 <br> leaf vegetables cooked 0.54 <br> onions, garlic 0.52 <br> cabbage 0.43 <br> fish 0.42 |

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|  |  | crustaceans mullocs 0.42 <br> mushrooms 0.37 <br> seed oils 0.33 <br> tomatoes cooked 0.32 <br> fresh fruit (non-citrus) 0.32 <br> nuts and seeds 0.30 |
| :---: | :---: | :---: |
| (Paradis, Perusse \& Vohl, 2006) <br> Different Gender <br> Men Women | red meats 0.780 .80 <br> butter 0.62 <br> poultry 0.590 .63 <br> high-fat dairy products 0.56 <br> processed meats 0.550 .68 <br> potatoes other than french fried 0.47 <br> refined grains0.46 0.38 <br> condiments 0.44 <br> french fries 0.44 <br> mayonnaise 0.36 <br> desserts0.33 0.30 | vegetables 0.760 .65 <br> fruits 0.680 .73 <br> non-hydrogenated fat 0.62 <br> fish and other seafood 0.33 <br> wine 0.30 <br> nuts0.48 <br> legumes 0.47-0.41 <br> whole grains 0.41 <br> organ meats 0.37 <br> fruit juices -0.33 |
| (Ronco et al., 2006) |  | fried meat 0.81 <br> barbecue 0.66 <br> processed meat 0.47 |
| (Cottet et al., 2005) (women only) | poultry 0.47 <br> fish and crustaceans 0.32 <br> high-fat delicatessen 0.37 <br> vegetable fat 0.49 <br> nuts 0.48 <br> legumes 0.30 <br> potatoes 0.67 <br> refined bread and cereals 0.36 <br> milk 0.30 <br> sodas 0.35 <br> beer 0.39 <br> condiments 0.43 |  |
| (Marchioni et al., 2005) |  | meat 0.617 <br> vegetables 0.817 <br> fruits 0.651 |

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| (Montonen et al., 2005) |  |  |
| :---: | :---: | :---: |
| (Nkondjock et al., 2005) | ```processed meats 0.52 sweets and desserts 0.49 refined grains 0.48 potatoes 0.43 processed fish 0.41 organ meats 0.39 soft drinks 0.36 legumes and legume products 0.36 snacks 0.34 margarine 0.33 nuts 0.32``` |  |
| (Perrin et al., 2005) | sugar and sweets 0.70 <br> grains 0.54 <br> butter 0.50 <br> added fats 0.44 <br> eggs 0.43 <br> dairy products 0.37 | fruit 0.56 <br> vegetables 0.54 <br> olive oil 0.48 <br> fish 0.32 <br> alcohol 0.44 <br> high-fat meat 0.50 <br> potatoes 0.55 <br> dairy products 0.34 |
| (Rashidkhani et al., 2005) | sweets 0.56 <br> processed 0.55 <br> refined grains 0.54 <br> added fat 0.51 <br> high-fat dairy 0.49 <br> fried potatoes 0.41 <br> soft drinks 0.4 <br> meat beef 0.4 <br> cooked 0.33 |  |

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| (Khani et al., 2004) | sweets 0.56 <br> processed meat 0.55 <br> refined grains 0.54 <br> margarine 0.51 <br> high-fat dairy 0.49 <br> fried potatoes 0.41 <br> soda 0.40 <br> meat 0.40 <br> cooked potato 0.33 |
| :---: | :---: |
| (Uusitalo et al., 2005) <br> Different time points 199819921998 | rice -0.43-0.34-0.11 <br> bread 0.270 .200 .12 <br> pulses -0.30 -0.10 0.00 <br> poultry 0.240 .240 .40 <br> processed meat 0.150 .140 .40 <br> fresh/frozen fish 0.140 .250 .40 <br> butter -0.18-0.10-0.02 <br> margarine 0.240 .080 .11 <br> whole milk -0.16-0.39-0.09 <br> skimmed/low-fat milk $0.33,0.410 .17$ |
| (Kim et al., 2004) | butter 0.400 .37 <br> mayonnaise 0.370 .36 <br> cheese 0.480 .38 <br> beef 0.540 .45 <br> pork 0.390 .48 <br> poultry 0.400 .45 <br> bacon 0.490 .55 <br> liver 0.460 .38 <br> soda beverages 0.350 .42 <br> fruit juice 0.390 .40 <br> vegetable juice 0.380 .32 <br> instant noodles 0.340 .31 <br> coffee 0.210 .26 <br> black tea 0.250 .24 |

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## Table E . Descriptive Statistics of our systematic review

| Justifications for the use of PCA vs. Single Food or Nutrient Analysis | Number / total number of papers | Percentage (\%) unless otherwise stated |
| :---: | :---: | :---: |
| - Interactive, antagonistic and synergistic effects of foods consumed in combination. | 25/163 | 15.3 |
| - Additive effects of foods consumed in combination which are too small to detect when they are examined separately | 14/163 | 8.5 |
| - Confounding and mulitcollinearity from lifestyle factors and dietary exposures | 52/163 | 31.9 |
| - Multiple testing problems | 7/163 | 4.2 |
| - Public health recommendations | 14/163 | 8.5 |
| - Complexity of diet | 25/163 | 15.3 |
| - Better evaluation / representation of overall diet | 30/163 | 18.4 |
| Study Design |  |  |
| - Case-control | 31/163 | 17.1 |
| - Cohort | 31/163 | 19.0 |
| - Cross-sectional | 99/163 | 64.0 |
| - Other | 4/163 | 2.4 |
| Dietary assessment instrument |  |  |
| - Food Frequency Questionnaire | 152/163 | 93.2 |
| - 24/48 hour recall | 4/163 | 2.4 |
| - Dietary records | 3/163 | 1.8 |
| - Diet history questionnaire | 3/163 | 1.8 |
| - number of food items in each instrument | 163 | median value:92 ( IQR:21-204) |
| - number of food groups in each instrument | 163 | median value:38 (IQR:19-69) |
| - Scale of Food Frequency Questionnaire | 163 | median value:7 (IQR: 5-10) |
| Preparation of data before entering the PCA |  |  |
| - Conversion of food frequency data to grams/d or grams/week | 22/163 | 13.4 |
| - Food items collapsed to food groups | 42/163 | 28.2 |
| - Standardisation of food intake variables | 6/163 | 3.6 |

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| - Food intake variables adjusted for energy intake by the residual method | 5/163 | 3.0 |
| :---: | :---: | :---: |
| Labeling the dietary patterns in each study |  |  |
| - principal component loading and correlation coefficient cut off points in each study | 163 | median value:0.3 (IQR: 0.3-0.4) |
| Method of rotation in each study |  |  |
| - Orthogonal/ Varimax | 85/163 | 52.1 |
| - Oblique / Promax | 5/163 | 3.0 |
| Identifying the number of dietary patterns in each study |  |  |
| - Scree Plot | 81/163 | 49.1 |
| - cut-off point s for eigenvalues | 163 | median value:1.6; (IQR: 1-2) |
| - dietary patterns interpretability | 70/163 | 42.9 |
| - Van der Voet's test | 3/163 | 1.8 |
| Validation methods of retained dietary patterns | 40/163 | 24\% |
| - Randomly split sample in each study | 8/163 | 4.9 |
| - Cronbach's alpha | 5/163 | 3.0 |
| - Deriving dietary patterns separately for women and men | 3/163 | 1.8 |
| - Barlet's test of sphericity | 1/163 | 0.6 |
| - Kaiser - Meyen-Ollkin test | 2/163 | 1.2 |
| - Identifying simplified dietary patterns | 3/163 | 1.8 |
| - Use of Confirmatory/maximum likelihood factor analysis | 8/163 | 4.9 |
| - $\varphi$ coefficient for testing inter-correlation | 2/163 | 1.2 |
| - Stricter cut-off points for energy intake | 1/163 | 0.6 |
| - Different method of rotation | 2/163 | 1.2 |
| - Pearson Correlation coefficient in different time points | 5/163 | 3.0 |
| Number of dietary patterns | 163 | median: 3 (IQR: 2-4) |
| Percentage of total variance of original food items being explained by the dietary patterns in each study. | 163 | median: 24(IQR: 19.9-31.3) |
| Number of studies that examine associations with health outcomes | 120/163 | 73.6 |
| Number of studies that examine associations with socio-economic characteristics. | 60/163 | 36.8 |
| Number of studies which didn't associated Principal Components with any health outcome or socio-economic factor. | 29/163 | 17.7 |

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## Appendices

## I. Data Abstraction questions for the systematic review

The relevant information was abstracted from eligible studies and inserted into an Excel data collection form with prepared fields. Data collected included:

1. Title
2. Author
3. Justifications for the use of PCA instead of a single food or nutrient analysis
(1. Interactive and synergistic effects of foods consumed in combination. 2. Additive Effects of foods consumed in combination 3. Mulitcollinearity and confounding between dietary and lifestyle exposures 4. Multiple testing problems 5. Public health recommendation 6. Complexity of diet, 7.Better evaluation / representation of overall diet)
4. Details of the Study
(1.Population being used in each study 2. sample size being used in each study)
5. Country/Countries
6. Health outcome being investigated
7. Details of dietary assessment instrument being used
(1.Instrument Definition a.FFQ b.24/48- hour recall c. dietary records d. diet history questionnaire (EPIC/DHQ)
8. Number of food items of the instrument 3. Number of food groups of the instrument 4. Scale of FFQ)
9. Preparing the data before entering the PCA procedure
(1.Convertion to grams/d or grams /week 2. Standardisation of food intake variabless3. Food intake variables adjusted for energy intake by the residual method 4. Box-Cox transformation method)
10. Criteria for labelling and identifying the number of dietary patterns
(1. Cut off points 2. Method of rotation a. Varimax b. Oblique (premix) 3. Scree Plot 4. Eigenvalues 5. Principal component interpretability 6. Van der Voet's )
11. Application of PCA
(1. Number of dietary patterns 2.Percentage of total variance being explained by the dietary patterns )
12. Foods included in dietary pattern 1 and label of the dietary pattern
13. Foods included in dietary pattern 2 and label of the dietary pattern
14. Foods included in dietary pattern 3 and label of the dietary pattern
15. Foods included in dietary pattern 4 and label of the dietary pattern
16. Foods included in dietary pattern 5 and label of the dietary pattern
17. Foods included in dietary pattern 6-10 and label of the dietary pattern
18. Associations of dietary patterns with health outcome
19. Associations of dietary patterns with socio-demographic variables
20. Validation methods (1.Two random split sample 2. Cronbach's alpha 3. Deriving patterns separately for women and men4. Barlet's test of sphericity5. Kaiser - Meyen-Ollkin test 6. Simplified Dietary patterns 7. Confirmatory/maximum likelihood factor analysis 8. $\varphi$ coefficient for testing inter-correlation 9. Stricter cut-off points for energy intake 10. Different method of rotation 11. Pearson Correlation different time points
21. Conclusions
22. Personal Criticism

## II. Code used to develop and analyze simulation experiment

```
*** Code used to develop and analyze simulation data ***
*** Estimating power of ESFA and PCA ***
program drop mybootun
program mybootun , rclass
display " Monte Carlo '
display " ------------------------------- '
display " Varying Effect Size and Principal Components "
display " ------------------------------- "
display " effect size is `4' , constant=`3'"
display " ------------------------------------
display " "" odds ratio is " exp(`4') ", cut off value=`7'"
display " ---------------------------------
display " `6' random selected foods , `5' principal components , sample size= `2' , number of replications=`1', food items = `9'
"
display " -----------------------------------
clear
quietly set memory 64m
set matsize 500
quietly postutil clear
local sum = 0
local sum1 =0
forvalues i=1(1)`1' {
```


## Appendices

quietly use "C:\Phd\Phd\Phd\Phd\numberl.dta", replace
quietly sample ‘6', count
forvalues j=1(1)`6' \{ local k`j' = number[`j'] \} quietly use "C:\Phd\echrsbootstm.dta" , replace quietly bsample '2' local conf " port`k1' "
local v=(`6'/2)
forvalues $w=2(1)$ 'v' \{

```
local conf "`conf' + port`k`w''"
```

\}
local $z=` v^{\prime}+1$
forvalues $w=z^{\prime}(1){ }^{\prime} 6^{\prime}$ \{
local conf "`conf' - port`k`w''" \} quietly gen \(p=1 /\left(1+\exp \left(-\left(3^{\prime}+{ }^{\prime} \mathbf{4}^{\prime *}\left(` \operatorname{conf}{ }^{\prime}\right)\right)\right)\right)\)
gen uniform=uniform()
quietly gen disease $=0$
quietly replace disease $=1$ if uniform $<=p$
quietly pca port6-port79, comp(`5') quietly rotate, varimax quietly predict pc1-pc`5'
postfile multipc1 coef pvalue using "C:\Phd\Phd\Phd\Phd\multipc2.dta" , replace
forvalues i=1(1)'5' \{
local b`i'=_b[pc`i']

## Appendices

```
local z`i'=_b[pc`i']/_se[pc`i']
local pval`i'= 2*normal(-abs(`z`i''))
post multipc1 (`b`i'') (`pval`i'')
}
```

postclose multipc1
postfile multisf1 coef pvalue using "C:\Phd\Phd\Phd\Phd\multisf2.dta" ,replace foreach var of varlist port6-port79 \{

```
quietly xi:logit disease `var' alkcal
    if e(df_m)> 1 {
                                    local bs=_b[`var']
                                    local zs=_b[`var']/_se[`var']
                                    local pvals=2*normal(-abs(`zs'))
                                    post multisf1 (`bs') (`pvals')
                                    }
```

postclose multisf1
quietly use "C:\Phd\Phd\Phd\Phd\multipc2.dta" ,replace
quietly multproc, pvalue(pvalue)
if r(nreject) > 0 \{
local sum = ‘sum' + 1
\}
quietly use "C:\Phd\Phd\Phd\Phd\multisf2.dta" , replace
quietly multproc, pvalue (pvalue)
if $r($ nreject $)>0$ \{
local sum1 = 'sum1'+
\}
local prop = `sum'/ `1'
local prop1 = 'sum1'/ '1'
display " proportion of any statistical effect is "prop' " display " proportion of any statistical effect is 'prop1'
end
local i=log(1.5)
mybootun $100001200-1.74$ 'i' 2100.379
local i=log(1.5)
mybootun 10000300 -1.74 ‘i' 2100.379
local $i=\log (1.5)$
mybootun $100004800-1.74$ 'i' 2100.379
local i=log(1.5)
mybootun $100001200-1.74$ 'i' $510 \quad 0.379$
local $i=\log (1.5)$
mybootun $10000300-1.74$ 'i' 5100.379
local $i=\log (1.5)$
mybootun $100004800-1.74$ 'i' 5100.37
local i=log(1.5)
mybootun $10000 \quad 1200-1.74$ 'i' $1010 \quad 0.3 \quad 79$
local $i=\log (1.5)$
mybootun $10000300-1.74$ 'i' $1010 \quad 0.379$
local $i=\log (1.5)$
mybootun $100004800-1.74$ 'i' $1010 \quad 0.379$
*** Estimating Power and FDR of ESFA ***
program drop mybootunmix
program mybootunmix , rclass
display " Monte Carlo"
display " ------------------------------ "

## Appendices

display " Varying Effect Size "
display " $\qquad$ "
display " effect size is '4' , constant=`3' , Simes procedure" display " --------------------------------" display " FDR at " `8'*100 "\%" " level, " " odds ratio is " exp(`4') ", cut off value=`7'"
display " ------------------------------- "
display " `6' random selected foods , '5' principal components , sample size= `2' , number of replications=`1' , food items=`9'
display " $\qquad$ - "
clear
quietly set memory 64 m
set matsize 500
quietly postutil clear
postfile power1 power1 using "C:\Phd\power1.dta" , every (1)replace postfile fdrl fdr1 using "C:\Phd\fdr1.dta" ,every(1) replace
forvalues $i=1(1)^{\prime} 1$ ' \{
local S1=0
local R1=0
quietly use "C:\Phd\number1.dta" ,replace
quietly sample '6', count
forvalues j=1(1)'6' \{

$$
\text { local } \mathrm{k}^{\prime} j^{\prime}=\text { number }\left[` j^{\prime}\right]
$$

\}
quietly use "C:\Phd\echrsbootstm.dta" , replace
quietly bsample '2'
local conf " port`k1' " local \(\mathrm{v}=(` 6\) '/2)

## Appendices

```
forvalues w=2(1)`v' {
    local conf "`conf' + port`k`w''"
    }
local z=`v'+1
forvalues w=`z'(1)`6' {
    local conf "`conf' + port`k`w''"
}
quietly gen p=1 /(1+exp(-(`3'+`4'*(`conf'))))
gen uniform=uniform()
quietly gen disease = 0
quietly replace disease = 1 if uniform <= p
postfile multisf1 coef pvalue using "C:\Phd\multisf1.dta" ,replace
forvalues k=6(1)79 {
quietly xi:logit disease port`k' alkcal
if e(df_m) > 1 {
    local bs=_b[port`k']
    local zs=_b[port`k']/_se[port`k']
    local pvals=2*normal(-abs(`zs'))
    post multisf1 (`bs') (`pvals')
    }
else {
local bs=0
```


## Appendices

```
    local pvals=1
    post multisf1 (`bs') (`pvals')
```

\}
\}
postclose multisf1
quietly use "C:\Phd\multisf1.dta" ,replace
quietly multproc, pvalue(pvalue) method(simes) reject(dummy1) puncor(`8')
quietly tab dummy
if $r($ nreject $) \quad$ > 0

```
                                    forvalues i=1(1)`6' {
```

                                    local v=(`k`i''-5)
                                    if dummy1[`v']==1 \{
                                    local S1 = `S1'+1
    \}
local power=(`S1'/`6')*100 post power1 ('power')
\}
else \{
local power=0
post power1 ('power')
\}

## Appendices

```
if r(nreject) > 0 {
    forvalues i=1(1)74 {
                                    if dummy1[`i']==1 {
                                    local R1 = `R1'+1
                                    }
    if `R1'==0 {
        local fdr=0
        post fdr1 (`fdr')
        }
    else {
        local fdr=((`R1' -`S1')/`R1')*100
        post fdr1 (`fdr')
            }
                        }
        else {
            local fdr=0
            post fdr1 (`fdr')
    }
}
postclose power1
postclose fdr1
use "C:\Phd\power1.dta" ,replace
tab power1
sort power1
sum power1
local seboot=r(sd)/sqrt(`1')
display "standard error is " `seboot'
```


## Appendices

```
local lowerp=`1'*(0.05/2)
local upperp=`1'*(1-(0.05/2))
display " 95% CI (" powerl[int(`lowerp')] ","power1[int(`upperp')] ")"
use "C:\Phd\fdr1.dta" ,replace
sort fdr1
sum fdr1
local seboot= r(sd)/sqrt(`1')
display "standard error is " `seboot'
local lowerp=`1'*(0.05/2)
local upperp=`1'*(1-(0.05/2))
display " 95% CI (" fdr1[int(`lowerp')] ","fdr1[int(`upperp')] ")"
end
local i=log(1.5)
mybootunmix 10 1200 -1.74 `i' 5 10 0.3 0.2 79
local i=log(1.5)
mybootunmix 10000 2400 -1.74 `i' 5 10 0.3 0.2 79
local i=log(1.5)
mybootunmix 10000 300 -1.74 `i' 5 10 0.3 0.2 79
local i=log(1.5)
mybootunmix 10000 600 -1.74 `i' 5 10 0.3 0.2 79
local i=log(1.5)
mybootunmix 10000 4800 -1.74 `i' 5 10 0.3 0.2 79
*** Estimating Power and FDR of PCA ***
program drop mybootpcaneg
program mybootpcaneg, rclass
display " Monte Carlo "
display "
display " Varying Sample size "
display "
display " effect size is `4' , constant=`3' , Simes procedure"
display " ----------------------------------"
display " FDR at " `8'*100 "%" " level , " " odds ratio is " exp(`4') ", cut off value=`7'"
```


## Appendices

```
display " ------------------------------- "
display " `6' random selected foods , `5' principal components , sample size= `2' , number of replications=`1' , food items=`9'
display "
```

$\qquad$

``` - "
clear
quietly set memory 400 m
set matsize 500
quietly postutil clear
postfile power power using "C:\Phd\Phd\Phd\Phd\power.dta" , every (1) replace
postfile fdr fdr using "C:\Phd\Phd\Phd\Phd\fdr.dta" ,every (1) replace
forvalues i=1(1) '1' \{
```

```
local S = 0
```

local S = 0
local R =0
local R =0
quietly use "C:\Phd\Phd\Phd\Phd\number1.dta" ,replace
quietly sample `6', count forvalues j=1(1)`6' {
local k`j' = number[`j']
}
quietly use "C:\Phd\echrsbootstm.dta" ,replace
quietly bsample `2'                                     local conf " port`k1' "
local v=(`6'/2) forvalues w=2(1)`v' {
local conf "`conf' + port`k`w''"
}

```

\section*{Appendices}
```

local z=`v'+1 forvalues w=`z'(1)`6'                                     local conf "`conf' - port`k`w''"
}
quietly gen p=1 /(1+exp(-(`3'+`4'*(`conf')))) gen uniform=uniform() quietly gen disease = 0 quietly replace disease = 1 if uniform <= p quietly pca port6-port79, comp(`5')
quietly rotate, varimax
quietly predict pc1-pc`5' quietly corr port`k1' port`k2' port`k3' port`k4' port`k5' port`k6' port`k7' port`k8' port`k9' port`k10' pc1-pc`5'
quietly matrix A=r(C)
quietly corr port6-port79 pc1-pc`5' quietly matrix A1=r(C) postfile multipc1 coef pvalue using "C:\Phd\Phd\Phd\Phd\multipc1.dta" ,replace forvalues i=1(1)`5' {
quietly xi: logit disease pc`i' alkcal     local b`i'=_b[pc`i']     local z`i'=_b[pc`i']/_se[pc`i']
local pval`i'= 2*normal(-abs(`z`i'''))     post multipc1 (`b`i'') (`pval`i'')
}
postclose multipc1

```

\section*{Appendices}
```

quietly use "C:\Phd\Phd\Phd\Phd\multipc1.dta" ,replace
quietly multproc, pvalue(pvalue) reject(dummy) pcor(0.05)
if r(nreject)> 0 {
forvalues j=1(1)`6' {                                     local w=0                                     forvalues i=1(1)`5' {
if abs(A[`6'+`i',`j']) >=`7' \& dummy[`i']==1 & `w'<1 {
local S = 'S'+ 1
}
}
local power=(`S'/`6')*100
post power (`power')                                     } else {     local S=0     local power=0     post power (`power')
}
if r(nreject)> 0 {
forvalues j=1(1)74 {
local l=0

```
```

                                    forvalues i=1(1)`5' {
    if abs(A1[74 +`i',`j']) >= `7' & dummy[`i']==1 \& `l'<1 {                                     local R = `R'+ 1

```
            local l=1
                if \({ }^{\prime} \mathrm{R}^{\prime}==0\)
                                    local fdr=0
                                    post fdr (`fdr')
            else \{
            local fdr=((`R' -`S')/`R')*100
            post fdr ('fdr')
            \}
    else \{
        local fdr=0
        post fdr (`fdr')
    \}
\(\}\)
postclose power
postclose fdr
use "C:\Phd\Phd\Phd\Phd\power.dta" ,replace
sort power
sum power
local seboot= r(sd)/sqrt(`1')
display "standard error is " sseboot'
local lowerp=`1'*(0.05/2)
local upperp=`1'*((1-0.05)/2)
display " 95\% CI (" power[int(`lowerp')] ", "power[int(`upperp')] ")"
use "C:\Phd\Phd\Phd\Phd\fdr.dta", replace
sort fdr
sum fdr
local seboot= \(r(s d) / \operatorname{sqrt}\left({ }^{\prime} 1^{\prime}\right)\)
display "standard error is " 'seboot'
local lowerp=`1'*(0.05/2)
local upperp=`1'*(1-(0.05/2))
display " 95\% CI (" fdr[int(`lowerp')] ","fdr[int(`upperp')] ")"
end
local \(i=\log (1.5)\)
mybootpcaneg \(101200-1.74\) 'i' 2100.30 .2
local i=log(1.5)
mybootpcaneg \(10000300-1.74\) 'i' 2100.30 .2
local i=log(1.5)
mybootpcaneg \(100004800-1.74\) `i' 2100.30 .2
local i=log(1.5)
mybootpcaneg 100001200 -1.74 'i' 5100.30 .2
local i=log(1.5)
mybootpcaneg 10000300 -1.74 ‘i' 5100.30 .2
local i=log(1.5)
mybootpcaneg \(100004800-1.74\) 'i' \(510 \quad 0.30 .2\)

\section*{Appendices}
```

local i=log(1.5)
mybootpcaneg 10000 1200 -1.74 i' 10 10 0.3 0.2
local i=log(1.5)
mybootpcaneg 10000 300 -1.74 'i' 10 10 0.3 0.2
local i=log(1.5)
mybootpcaneg 10000 4800 -1.74 'i' 10 10 0.3 0.2
*** Estimating power and FDR of ESFA adjusting for propensity scores ***
program drop mybootadjneg
program mybootadjneg , rclass
display " Monte Carlo"
display "
display " Varying Effect Size "
display "
display " effect size is `4' , constant=`3' , Simes procedure"
display " ------------------------------"
display " FDR at " ` ''*100 "%" " level , " " odds ratio is " exp(`4') ", cut off value=`7'" display " ------------------------------------ display " `6' random selected foods , `5' principal components , sample size= `2' , number of replications=`1', food items= `9'
display '

```
\(\qquad\)
``` - "
clear
quietly set memory 64 m
set matsize 500
quietly postutil clear
postfile power1 power1 using "C:\Phd\Phd\Phd\Phd\power1.dta" , every (1)replace
postfile fdr1 fdr1 using "C:\Phd\Phd\Phd\Phd\fdr1.dta" ,every(1) replace
```

```
forvalues l=6(1)`9' {
local conf " "
forvalues j=6(1)`9' {
if `j'!=`l' {
local conf "`conf' port`j'"
```

```
                                    }
                                    local ps`l' "`conf'"
                                    }
                                    }
forvalues i=1(1)`1' {
    local S1=0
local R1=0
quietly use "C:\Phd\Phd\Phd\Phd\number1.dta" ,replace
quietly sample `6', count
forvalues j=1(1)`6' {
                                    local k`j' = number[`j']
                                    }
quietly use "C:\Phd\Phd\Phd\Phd\echrsbootstm.dta" ,replace
quietly bsample `2'
                                    local conf "port`k1'"
local v=(`6'/2)
forvalues w=2(1)`v' {
    local conf "`conf' + port`k`w''"
    }
local z=`v'+1
```


## Appendices

forvalues w=`z'(1) '6' \{ local conf "`conf' - port`k`w''"
\}
quietly gen $p=1 /\left(1+\exp \left(-\left(3^{\prime}+{ }^{\prime} 4^{\prime} *(` \operatorname{conf})\right)\right)\right)$
gen uniform=uniform()
quietly gen disease $=0$
quietly replace disease $=1$ if uniform $<=p$
postfile multisf1 coef pvalue using "C:\Phd\Phd\Phd\Phd\multisf1.dta" ,replace
forvalues k=6(1)'9' \{

```
quietly regress port`k' `ps`k''
quietly predict ps`k'
quietly xi:logit disease port`k' ps`k' alkcal
local port " "
if e(df_m) > 2 {
```

local bs=_b[port'k']
local zs=_b[port`k']/_se[port`k']
local pvals=2*normal(-abs(`zs')) post multisf1 (`bs') (`pvals')
\}
else \{

## Appendices

```
    local k=1
    local bs=0
    local pvals=1
    post multisf1 (`bs') (`pvals')
```

\}
\}
postclose multisf1
quietly use "C:\Phd\Phd\Phd\Phd\multisf1.dta" ,replace
quietly multproc, pvalue(pvalue) method(simes) reject(dummy1) puncor(`8') if \(r(\) nreject \()>0\) forvalues i=1(1)'6' \{                                     local v=(`k`i''-5)                                     if dummy1['v']==1 \{                                     local \(\mathrm{S} 1=\) 'S1'+1                                     \}                                     local power=(`S1'/`6')*100                                     post power1 (`power')
\}
else \{
local power=0

```
    post power1 (`power')
    }
if r(nreject) > 0 {
            forvalues i=1(1)74 {
                                    if dummy1[`i']==1 {
                                    local R1 = `R1'+1
                                    }
                                    if `R1'==0 {
                                    local fdr=0
                                    post fdr1 (`fdr')
                                    else {
                                    local fdr=((`R1' -`S1')/`R1')*100
                                    post fdr1 (`fdr')
            }
                }
    else {
        local fdr=0
    post fdr1 (`fdr')
    }
```


## Appendices

\}
postclose powerl
postclose fdrl
use "C:\Phd\Phd\Phd\Phd\power1.dta" , replace
tab power1
sort power1
sum power1
local seboot= r(sd)/sqrt(`1') display "standard error is " ‘seboot' local lowerp=`1'*(0.05/2)
local upperp=`1'*(1-(0.05/2)) display " 95\% CI (" powerl[int(`lowerp')] ","power1[int(`upperp')] ")" use "C:\Phd\Phd\Phd\Phd\fdr1.dta", replace sort fdr1 sum fdrl local seboot= r(sd)/sqrt(`1')
display "standard error is " 'seboot'
local lowerp=`1'*(0.05/2) local upperp=`1'*(1-(0.05/2))
display " 95\% CI (" fdr1[int(`lowerp')] ","fdr1[int(`upperp')] ")"
end
local i=log(1.5)
mybootadjneg 10300 -1.74 'i' 540.30 .279
local i=log(1.5)
mybootadjneg $100001200-1.74$ 'i' 540.30 .279
local i=log(1.5)
mybootadjneg 100004800 -1.74 `i' 540.30 .279
*** Estimating Power and FDR of ESFA adjusted for 5 principal components ***
program drop mybootadjpcneg
program mybootadjpcneg , rclass
display " Bootstrap"
display " ------------------------------- "

## Appendices

```
display " Varying Effect Size "
display '
display " Modelling Disease with main effects only, effect size is `4' , constant=`3' , Simes procedure"
display " ----------------------------------"
display " FDR at " `8'*100 "%" " level , " " odds ratio is " exp(`4') ", cut off value=`7'"
display " ---------------------------------"
display " `6' random selected foods , `5' principal components , sample size= `2' , number of replications=`1' , food items = `9'
display '
```

$\qquad$

``` - "
clear
quietly set memory 64 m
set matsize 500
quietly postutil clear
postfile power1 power1 using "C:\Phd\power1.dta" , every (1)replace
postfile fdrl fdrl using "C:\Phd\fdr1.dta" ,every(1) replace
local f=`9'+ 5
forvalues i=1(1)`1' \{
local S1=0
local R1=0
quietly use "C:\Phd\number1.dta" ,replace
quietly sample '6', count
forvalues j=1(1)`6' \{
\[
\text { local } \mathrm{k}^{\prime} j^{\prime}=\text { number }\left[` j^{\prime}\right]
\]
\(\}\)
quietly use "C:\Phd\echrsbootstm.dta" ,replace quietly bsample '2'
local conf " port`k1'"
local v=(`6'/2)
```

```
forvalues w=2(1)`v' {
    local conf "`conf' + port`k`w''"
    }
local z=`v'+1
forvalues w=`z'(1)`6' {
    local conf "`conf' + port`k`w''"
    }
quietly gen p=1 /(1+exp(-(`3'-`4'*(`conf'))))
gen uniform=uniform()
quietly gen disease = 0
quietly replace disease = 1 if uniform <= p
quietly pca port6-port`f', comp(`5')
quietly rotate, varimax
quietly predict pcl-pc`5'
local conf1 " "
forvalues j=1(1)`5' {
```

```
local conf1 "`conf1' pc`j'"
```

local conf1 "`conf1' pc`j'"
}

```

\section*{Appendices}
postfile multisf1 coef pvalue using "C:\Phd\multisf1.dta", replace
forvalues \(k=6(1)\) 'f' \{
```

quietly xi:logit disease port`k' `conf1' alkcal
if e(df_m) > (`5'+1) {

```
    local bs=_b[port`k']
    local zs=b[port`k']/ se[port`k']
    local pvals=2*normal(-abs(`zs'))
    post multisf1 (`bs') (`pvals')
            \}
else \{
    local bs=0
    local pvals=1
    post multisf1 (`bs') (`pvals')
    \}
postclose multisf1
quietly use "C:\Phd\multisf1.dta" , replace
quietly multproc, pvalue(pvalue) method(simes) reject(dummy1) puncor(`8') quietly tab dummy
if r(nreject) > 0 \{
forvalues i=1(1)`6' \{
if dummy1[`v']==1 \{
local \(\mathrm{S} 1=\) 'S1'+1
\}
local power=(`S1'/`6')*100
post power1 (`power')
\}
else \{
local power=0
post power1 ('power')
if r(nreject) > 0 \{
forvalues i=1(1)'9' \{
if dummy1[`i']==1
local \(\mathrm{R} 1=\) 'R1'+1
\}
if 'R1'==0 \{
local fdr=0 post fdr1 (`fdr')
else \{

\section*{Appendices}
```

local fdr=(`R1' -`S1')/`R1' post fdr1 (`fdr')
}
}
else {
local fdr=0
post fdr1 (`fdr') } } postclose powerl postclose fdrl use "C:\Phd\power1.dta" ,replace tab power1 sort power1 sum power1 local seboot= \(r(s d) / \operatorname{sqrt}\left(` 1^{\prime}\right)\)
display "standard error is " `seboot' local lowerp=`1'*(0.05/2)
local upperp=`1'*(1-(0.05/2)) display " 95\% CI (" power1[int(`lowerp')] ","power1[int(`upperp')] ")" use "C:\Phd\fdr1.dta" ,replace sort fdrl sum fdr1 local seboot= r(sd)/sqrt(`1')
display "standard error is " ‘seboot'
local lowerp=`1'*(0.05/2) local upperp=`1'*(1-(0.05/2))
display " 95\% CI (" fdr1[int(`lowerp')] ","fdr1[int(`upperp')] ")"

```

\section*{Appendices}
end
local i=log(1.5)
mybootadjpeneg \(100001200-1.74\) 'i' 5100.30 .274
local i=log(1.5)
mybootadjpeneg \(10000300-1.74\) ‘i' \(510 \quad 0.30 .274\)
local i=log(1.5)
mybootadjpcneg \(100004800-1.74\) ‘i' 5100.30 .274
*** Estimating Power and FDR of ESFA adjusting for statistical significant foods ***
program drop mybootadjf
program mybootadjf , rclass
display " Monte Carlo"
display " \(\qquad\) "
display " Varying Effect Size "
display " -----------------------------------
display " Modelling Disease with main effects only, effect size is `4' , constant=`3' , Simes procedure"
display "
display " FDR at " ` 8 '*100 "\%" " level , " " odds ratio is " exp(`4') ", cut off value=`7"
display " ------------------------------ "
display " ‘6' random selected foods , '5' principal components , sample size= `2' , number of replications=`1' , food items = 9 ' "
display " \(\qquad\) - "
clear
quietly set memory 64 m
set matsize 500
quietly postutil clear
postfile power1 power1 using "C:\Phd\power1.dta" , every (1) replace
postfile fdr1 fdr1 using "C:\Phd\fdr1.dta" ,every (1) replace
local \(f=`{ }^{\prime}+5\)
forvalues i=1(1)'1' \{
```

local S1=0
local R1=0
quietly use "C:\Phd\number1.dta" ,replace
quietly sample `6', count forvalues j=1(1)`6' {
local k`j' = number[`j']
}
quietly use "C:\Phd\echrsbootstm.dta" ,replace
quietly bsample `2'                     local conf " port`k1'"
local v=(`6'/2) forvalues w=2(1)`v' {
local conf "`conf' - port`k`w''"     } local z=`v'+1
forvalues w=`z'(1)`6' {
local conf "`conf' + port`k`w''"
}

```

\section*{Appendices}
```

quietly gen p=1 /(1+exp(-(`3'+`4'*(`conf')))) gen uniform=uniform() quietly gen disease = 0 quietly replace disease = 1 if uniform <= p postfile multisf1 coef pvalue using "C:\Phd\multisf1.dta" ,replace forvalues k=6(1)`f' {
quietly xi:logit disease port`k' alkcal if e(df_m) > 1 {     local bs=_b[port`k']
local zs=__b[port`k']/_se[port`k']
local pvals=2*normal(-abs(`zs'))     post multisf1 (`bs') (`pvals')     }         else {     local bs=0     local pvals=1     post multisf1 (`bs') (`pvals')
}
}
postclose multisf1

```
```

preserve
quietly use "C:\Phd\multisf1.dta" ,replace
quietly multproc, pvalue(pvalue) method(simes) reject(dummy1) puncor(`8') quietly tab dummy1 local conf " " local w=0 if r(nreject) > 0 {     forvalues i=1(1)`9' {
if dummy1[`i']==1 {                                     local w=`i'+5
local conf "`conf' port`w'"
}
restore
postfile multisf1_a coef pvalue using "C:\Phd\multisf1_a.dta", replace
forvalues $k=6(1)$ f' \{
quietly xi:logit disease port`k' 'conf' alkcal local bs= b[port`k']
local $z s=$ _b[port'k']/_se[port $\left.{ }^{\prime}{ }^{\prime}\right]$
local pvals=2*normal(-abs(`zs'))

```
                                    \}

\section*{Appendices}
local bs=0
local pvals=1
post multisf1_a ('bs') ('pvals')
```

postclose multisf1_a
quietly use "C:\Phd\multisf1_a.dta" ,replace
quietly multproc, pvalue(pvalue) method(simes) reject(dummy1) puncor(`8') quietly tab dummy1 if r(nreject) > 0 {     forvalues i=1(1)`6' {
local v=(`k`i''-5)
if dummy1[`v']==1 {                                     local S1 = `S1'+1
}
}
local power=(`S1'/`6')*100
post power1 (`power')

```
```

}
else {
local power=0
post power1 (`power') if r(nreject) > 0 {     forvalues i=1(1)`9' {
if dummy1[`i']==1 {                                     local R1 = `R1'+1
}
if `R1'==0 {                         local fdr=0                                     post fdr1 (`fdr')
}
else {
local fdr=((`R1' -`S1')/`R1')*100                 post fdr1 (`fdr')
}
}
else {
local fdr=0
post fdr1 (`fdr')

```
```

dis `power dis `fdr'
postclose power1
postclose fdrl
use "C:\Phd\power1.dta" ,replace
sort power1
sum power1
local seboot= r(sd)/sqrt(`1') display "standard error is " `seboot'
local lowerp=`1'*(0.05/2) local upperp=`1'*((1-0.05)/2)
display " 95% CI (" power1[int(`lowerp')] ","power1[int(`upperp')] ")"
use "C:\Phd\fdr1.dta" ,replace
sort fdr1
sum fdr1
local seboot= r(sd)/sqrt(`1') display "standard error is " `seboot'
local lowerp=`1'*(0.05/2) local upperp=`1'*((1-0.05)/2)
display " 95% CI (" fdr1[int(`lowerp')] ","fdr1[int(`upperp')] ")"
end
local i=log(1.5)
mybootadjf 10000 1200 -1.74 `i' 5 10 0.3 0.2 74
local i=log(1.5)
mybootadjf 10000 300-1.74 'i' 5 10 0.3 0.2 74
local i=log(1.5)
mybootadjf 10000 4800 -1.74 'i' 5 10 0.3 0.2 74

```

\section*{Appendices}
*** Simulating disease from "Western" dietary pattern derived from the UK ECHRS II dataset instead of randomly selected foods (rest of the code remains the same ***
\begin{tabular}{ll} 
local & \(\mathrm{k} 1=13\) \\
local & \(\mathrm{k} 2=18\) \\
local & \(\mathrm{k} 3=21\) \\
local & \(\mathrm{k} 4=25\) \\
local & \(\mathrm{k} 5=30\) \\
local & \(\mathrm{k} 6=31\) \\
local & \(\mathrm{k} 7=37\) \\
local & \(\mathrm{k} 8=41\) \\
local & \(\mathrm{k} 9=52\) \\
local & \(\mathrm{k} 10=75\)
\end{tabular}
quietly gen p1=1 /(1+exp(-(`3'+`4'*(port`k1'+
port`k2'+port`k3'+port`k4'+port`k5'+port`k6'+port`k7'+port`k8'+port`k9'+port`k10'))))
gen uniform1=uniform()
quietly gen disease1 \(=0\)
quietly replace disease \(1=1\) if uniform1 \(<=p 1\)
*** Simulating disease from "Western" dietary pattern derived from the UK ECHRS dataset instead of randomly selected foods (rest of the code remains the same ***
\begin{tabular}{ll} 
local & \(k 1=6\) \\
local & \(k 2=28\) \\
local & \(k 3=31\) \\
local & \(k 4=34\) \\
local & \(k 5=112\) \\
local & \(k 6=113\) \\
local & \(k 7=114\) \\
local & \(k 8=115\) \\
local & \(k 9=119\) \\
local & \(k 10=121\) \\
local & \(k 11=122\)
\end{tabular}

\section*{Appendices}
\begin{tabular}{|c|}
\hline local k12=123 \\
\hline local k13=124 \\
\hline local k14=125 \\
\hline local k15=129 \\
\hline local k16=131 \\
\hline local k17=139 \\
\hline local k18=140 \\
\hline local k19=157 \\
\hline local k20=165 \\
\hline local k21=166 \\
\hline local k22=177 \\
\hline local k23=178 \\
\hline local k24=183 \\
\hline local k25=184 \\
\hline local k26=186 \\
\hline local k27=189 \\
\hline local k28=190 \\
\hline local k29=194 \\
\hline local k30=215 \\
\hline
\end{tabular}
quietly gen \(\mathrm{p} 2=1\) /(1+exp(-(`3'-`4'*(port`k1'+ port`k2'+
port`k3'+port`k4'+port`k5'+port`k6'+port`k7'+port`k8'+port`k9'+port`k10' +port`k11'+ port`k12'+ port`k13'+port`k14'+port`k15'+port`k16'+port`k17'+port`k18'+port`k19'+port`k20'+port`k21'+ port`k22'+

gen uniform2=uniform()
quietly gen disease2 = 0
quietly replace disease2 \(=1\) if uniform2 \(<=\) p2

\section*{III. Associations between specific food items that were declared significant from adjusted ESFA (FDR \(=5 \%\) ) and asthma: results of metaanalyses. OR; odds ratio.}




\begin{tabular}{|c|c|c|c|}
\hline \multicolumn{4}{|c|}{legumes} \\
\hline country & obs & & OR (95\% CI) \\
\hline Belgium & 143 & \[
\pm
\] & 1.29 (0.44, 3.75) \\
\hline Denmark & 352 & - & 0.82 (0.56, 1.19) \\
\hline Finland & 155 & \(+\) & 0.61 (0.23, 1.60) \\
\hline North Germany & 177 ¢ & - & 0.02 (0.00, 0.64) \\
\hline South Germany & 194 & - & 2.32 (0.75, 7.19) \\
\hline Holland & 211 & \(\square\) & 0.78 (0.39, 1.54) \\
\hline Portugal & 259 & - & 0.88 (0.72, 1.07) \\
\hline Poland & 210 & + & 1.14 (0.52, 2.51) \\
\hline UK & 171 & - & 0.35 (0.17, 0.71) \\
\hline Sweden & 1176 & - & 0.86 (0.73, 1.02) \\
\hline \multicolumn{3}{|l|}{Overall ( 1 -squared \(=41.5 \%, \mathrm{p}=0.081\) )} & 0.83 (0.68, 1.02) \\
\hline
\end{tabular}


couscous

turnip


\section*{IV. Associations between specific food items that were declared significant from adjusted ESFA (FDR \(=5 \%\) ) and chronic sinusitis: results of meta-analyses. OR; odds ratio.}

\begin{tabular}{|c|c|c|c|}
\hline country & obs & & OR (95\% CI) \\
\hline Denmark & 337 & ! & 1.07 (0.54, 2.14) \\
\hline Finland & 155 & \(\square\) & 1.10 (0.71, 1.72) \\
\hline North Germany & 172 & & 2.08 (0.67, 6.47) \\
\hline South Germany & 194 & \(\xrightarrow{+}\) & 0.88 (0.46, 1.69) \\
\hline Holland & 215 & \(\square\) & 1.91 (0.96, 3.80) \\
\hline Portugal & 259 & + & 1.10 (0.95, 1.28) \\
\hline Poland & 206 & \(\rightarrow\) & 0.83 (0.45, 1.51) \\
\hline UK & 171 & \(\square\) & 1.57 (0.86, 2.85) \\
\hline Sweden & 1176 & & 1.03 (0.06, 17.40) \\
\hline Belgium & 145 & & (Excluded) \\
\hline \multicolumn{2}{|l|}{Overall ( 1 -squared \(=0.0 \%, \mathrm{p}=0.618\) )} & 0 & 1.12 (0.99, 1.27) \\
\hline
\end{tabular}

\begin{tabular}{|c|c|c|}
\hline \multicolumn{3}{|c|}{pumpkin} \\
\hline country & obs & OR (95\% CI) \\
\hline Belgium & 145 & 1.03 (0.69, 1.54) \\
\hline Denmark & 337 & 1.14 (0.57, 2.28) \\
\hline Finland & 155 & 0.95 (0.26, 3.42) \\
\hline North Germany & 172 & 0.66 (0.11, 3.80) \\
\hline South Germany & 194 & 4.38 (1.09, 17.56) \\
\hline Holland & 215 & 1.73 (0.75, 3.97) \\
\hline Portugal & 259 & 1.10 (0.97, 1.26) \\
\hline Poland & 206 & 0.88 (0.52, 1.49) \\
\hline UK & 171 & 0.83 (0.02, 32.41) \\
\hline Sweden & 1176 & 0.97 (0.67, 1.41) \\
\hline Overall (l-square & \(d=0.0 \%, p=0.684)\) & 1.09 (0.97, 1.22) \\
\hline
\end{tabular}
V. Associations between specific food items that were declared significant from adjusted ESFA (FDR \(=5 \%\) ) and Allergic Rhinitis: results of meta-analyses. OR; odds ratio.





\begin{tabular}{|c|c|c|}
\hline country & obs & OR (95\% CI) \\
\hline Belgium & 146 & 1.28 (0.79, 2.09) \\
\hline Denmark & 353 & 1.15 (0.64, 2.08) \\
\hline Finland & 155 & 1.34 (0.80, 2.23) \\
\hline North Germany & 177 & 1.07 (0.09, 12.45) \\
\hline South Germany & 193 & 0.49 (0.15, 1.55) \\
\hline Holland & 215 & 0.28 (0.06, 1.43) \\
\hline Portugal & 259 & 1.02 (0.59, 1.78) \\
\hline Poland & 209 & 0.67 (0.32, 1.41) \\
\hline UK & 171 & 2.44 (0.29, 20.22) \\
\hline Sweden & 1173 & 1.12 (1.00, 1.27) \\
\hline \multicolumn{2}{|l|}{Overall ( 1 -squared \(=0.0 \%, \mathrm{p}=0.536\) )} & 1.11 (1.00, 1.24) \\
\hline
\end{tabular}
\begin{tabular}{|c|c|c|c|}
\hline \multicolumn{4}{|c|}{single cream} \\
\hline country & \multicolumn{2}{|l|}{obs} & OR (95\% CI) \\
\hline Belgium & 146 & & 3.67 (1.88, 7.16) \\
\hline Denmark & 353 & \(\rightarrow\) & 1.13 (0.98, 1.31) \\
\hline Finland & 155 & & 0.66 (0.16, 2.71) \\
\hline North Germany & 177 & & 1.15 (0.47, 2.81) \\
\hline South Germany & 193 & & 1.14 (0.60, 2.16) \\
\hline Holland & 215 & & 2.02 (0.90, 4.52) \\
\hline Portugal & 259 & & 0.84 (0.34, 2.07) \\
\hline Poland & 209 & \(\square\) & 1.25 (0.86, 1.82) \\
\hline UK & 171 & & 1.36 (0.58, 3.20) \\
\hline Sweden & 1173 & - & 1.07 (0.80, 1.43) \\
\hline \multicolumn{2}{|l|}{Overall ( 1 -squared \(=39.6 \%, \mathrm{p}=0.094\) )} & \(\bigcirc\) & 1.26 (1.03, 1.55) \\
\hline
\end{tabular}

\section*{VI. Associations between specific food items that were declared significant from adjusted ESFA (FDR \(=5 \%\) ) and eczema: results of metaanalyses. OR; odds ratio}




\begin{tabular}{|c|c|c|}
\hline \multicolumn{3}{|c|}{game} \\
\hline country & obs & OR (95\% CI) \\
\hline Belgium & 146 & 0.72 (0.19, 2.74) \\
\hline Denmark & 353 & 0.80 (0.29, 2.22) \\
\hline Finland & 155 & 1.03 (0.80, 1.32) \\
\hline North Germany & 177 & 0.31 (0.11, 0.87) \\
\hline South Germany & 194 & 0.56 (0.19, 1.70) \\
\hline Holland & 214 & 1.21 (0.24, 6.16) \\
\hline Portugal & 259 & 0.38 (0.12, 1.13) \\
\hline Poland & 209 & 5.22 (1.22, 22.30) \\
\hline UK & 168 & 18.05 (0.16, 2000.30) \\
\hline Sweden & 1176 & 0.94 (0.81, 1.09) \\
\hline Overall (l-square & = 42.4\%, \(\mathrm{p}=0.075\) ) & 0.88 (0.66, 1.17) \\
\hline
\end{tabular}

\begin{tabular}{|c|c|c|}
\hline country & obs & OR (95\% CI) \\
\hline Belgium & 146 & 0.37 (0.08, 1.74) \\
\hline Denmark & 353 & 0.80 (0.35, 1.82) \\
\hline Finland & 155 & 0.62 (0.18, 2.20) \\
\hline North Germany & 177 & 0.90 (0.46, 1.75) \\
\hline South Germany & 194 & 1.81 (0.98, 3.31) \\
\hline Holland & 214 & 1.20 (0.55, 2.65) \\
\hline Portugal & 259 & 0.10 (0.00, 5.27) \\
\hline Poland & 209 & 1.02 (0.82, 1.26) \\
\hline UK & 168 & 3.60 (0.61, 21.16) \\
\hline Sweden & 1176 & 1.08 (0.88, 1.33) \\
\hline \multicolumn{2}{|l|}{Overall ( 1 -squared \(=6.1 \%, p=0.385\) )} & 1.06 (0.91, 1.23) \\
\hline
\end{tabular}
VII. Associations between specific food items that were declared significant from adjusted ESFA (FDR \(=\mathbf{5 \%}\) ) and atopy: results of meta-analyses. OR; odds ratio.






\section*{IX. Associations between specific food items that were declared significant from adjusted ESFA (FDR \(=5 \%\) ) and respiratory and allergic outcomes: results of a random coefficient logistic model. OR; odds ratio.}
\begin{tabular}{|c|c|c|c|c|c|c|c|c|c|}
\hline \multicolumn{10}{|c|}{ESFA ( \({ }^{\text {nd }}\) step) \({ }^{\text {( }}\) FDR=5\% all P-values \(<0.001\)} \\
\hline \multicolumn{2}{|l|}{Asthma} & \multicolumn{2}{|l|}{Chronic Sinusitis} & \multicolumn{2}{|l|}{Allergic Rhinitis} & \multicolumn{2}{|l|}{Eczema} & \multicolumn{2}{|l|}{Atopy} \\
\hline food & OR & food & OR & food & OR & & OR & & OR \\
\hline kidney (black beans) & 0.85 & cabbage & 0.87 & peach & 0.77 & rhubarb & 0.84 & crisp fried cakes & 0.86 \\
\hline lemon & 0.85 & Viili yogurt like fermented milk & 1.11 & peanuts & 0.90 & \begin{tabular}{l}
crisp \\
fried \\
cakes
\end{tabular} & 0.84 & moussaka & 0.88 \\
\hline \begin{tabular}{l}
cured or \\
smoked \\
fatty fish
\end{tabular} & 0.85 & okra & 1.12 & tofu & 0.90 & bitter melon & 0.88 & beer & 1.10 \\
\hline cherries & 0.86 & pumpkin & 1.14 & mango & 1.10 & Greek style yogurt & 0.89 & thin biscuits (knackerb rod) & 1.13 \\
\hline legumes, any & 0.89 & & & vegetable oil & 1.17 & game, other & 0.92 & \begin{tabular}{l}
Hot cold \\
Roast \\
beef/ \\
boiled beef
\end{tabular} & 1.13 \\
\hline condensed milk & 0.88 & & & any smoked/ cured poultry & 1.09 & lentils & 0.93 & & \\
\hline thin biscuits (knackerbr od) & 1.09 & & & single cream & 1.24 & Sour cream & 1.15 & & \\
\hline turnip & 1.13 & & & & & & & & \\
\hline couscous & 1.17 & & & & & & & & \\
\hline
\end{tabular}
*Adjusted for age, sex, body mass index, smoking status and all the foods that were significant at the univariate analysis ( \(1^{\text {st }}\) step) ***Bold type indicates foods which were statistically significant associated with eczema across 3 or more of the different procedures being used

\section*{X. Description of mean and median intake (grams/per day) for each food item in our study}
\begin{tabular}{|c|c|c|c|}
\hline food item & mean (min-max) & median (p25-p75) & percentage of individuals with non zero consumption of the specific food item \\
\hline wholemeal bread & 145.90(0.00-350.00) & 137.50(25.00-175.00) & 88.82 \\
\hline white bread & 126.54(0.00-658.00) & 47.00(23.50-141.00) & 80.45 \\
\hline rye bread & 53.84(0.00-350.00) & 12.50(0.00-75.00) & 60.25 \\
\hline kneipp bread & 8.99 (0.00-560.00) & 0.00 (0.00-0.00) & 8.80 \\
\hline nan paratha bread & 9.27 (0.00-2100.00) & 0.00 (0.00-0.00) & 8.50 \\
\hline chappati bread & 3.78 (0.00-770.00) & 0.00 (0.00-0.00) & 6.61 \\
\hline wheat yeast rusks & 7.96 (0.00-420.00) & 0.00 (0.00-0.00) & 19.35 \\
\hline breakfast cereals & 98.57(0.00-560.00) & 40.00(0.00-220.00) & 66.54 \\
\hline couscous & 25.93(0.00-2100.00) & 0.00 (0.00-0.00) & 24.54 \\
\hline total pasta & 378.14(0.00-3220.00) & 230.00(115.00-690.00) & 89.44 \\
\hline refined pasta & 285.53(0.00-3220.00) & 230.00(115.00-230.00) & 84.03 \\
\hline wholemeal pasta & 105.44(0.00-3220.00) & 0.00 (0.00-115.00) & 40.34 \\
\hline filled pasta & 50.11(0.00-3220.00) & 0.00 (0.00-115.00) & 30.34 \\
\hline noodles & 55.45(0.00-3220.00) & 0.00 (0.00-115.00) & 30.43 \\
\hline any cakes & 113.57(0.00-840.00) & 60.00(30.00-180.00) & 85.96 \\
\hline cakes & \(46.25(0.00-840.00)\) & \(30.00(0.00-60.00)\) & 71.49 \\
\hline danish pastries & 44.21(0.00-1540.00) & 0.00 (0.00-55.00) & 45.06 \\
\hline sweet rolls & 5.53 (0.00-630.00) & 0.00 (0.00-0.00) & 11.53 \\
\hline muffins & 8.47 (0.00-840.00) & 0.00 (0.00-0.00) & 21.86 \\
\hline doughnuts/buns & 34.13(0.00-1050.00) & 0.00 (0.00-37.50) & 44.97 \\
\hline Puddings/ deserts & 47.96(0.00-2100.00) & 0.00 (0.00-75.00) & 37.70 \\
\hline custard cream & 13.24(0.00-840.00) & 0.00 (0.00-30.00) & 28.12 \\
\hline greek cakes & 1.75 (0.00-1400.00) & 0.00 (0.00-0.00) & 2.15 \\
\hline pancakes & 32.26(0.00-770.00) & 0.00 (0.00-55.00) & 49.53 \\
\hline Italian biscuits & 1.16 (0.00-280.00) & 0.00 (0.00-0.00 ) & 7.79 \\
\hline plain biscuits & 15.89(0.00-280.00) & 10.00(0.00-10.00) & 56.86 \\
\hline crisp fried cakes & 0.49 (0.00-140.00) & 0.00 (0.00-0.00) & 3.10 \\
\hline thin biscuits & 4.92 (0.00-56.00) & 2.00 (0.00-4.00) & 57.84 \\
\hline sweet biscuits & 3.86 (0.00-98.00) & 0.00 (0.00-3.50) & 43.34 \\
\hline rice & 137.32(0.00-1400.00) & 100.00(50.00-300.00) & 88.47 \\
\hline white rice & 116.88(0.00-1400.00) & 50.00(50.00-100.00) & 82.40 \\
\hline brown wholemeal pasta & 32.78(0.00-700.00) & 0.00 (0.00-50.00) & 35.13 \\
\hline rice noodles & 9.13 (0.00-1400.00) & 0.00 (0.00-0.00) & 11.50 \\
\hline table sugar & 15.49(0.00-70.00) & 2.50 (0.00-27.50) & 54.06 \\
\hline jam & 23.10(0.00-210.00) & 7.50 (0.00-15.00) & 67.77 \\
\hline marmelade & 10.13(0.00-210.00) & 0.00 (0.00-7.50) & 37.28 \\
\hline honey & 7.27 (0.00-112.00) & 0.00 (0.00-4.00) & 45.62 \\
\hline
\end{tabular}
\begin{tabular}{|c|c|c|c|}
\hline syrup spreads & 1.21 (0.00-70.00) & 0.00 (0.00-0.00 ) & 12.22 \\
\hline apple spreads & 1.24 (0.00-140.00) & 0.00 (0.00-0.00 ) & 8.67 \\
\hline total sweets & 37.97(0.00-280.00) & 20.00(0.00-60.00) & 73.97 \\
\hline Boiled sweets & 3.29 (0.00-70.00) & 0.00 (0.00-2.50) & 46.04 \\
\hline candies & 12.30(0.00-168.00) & 6.00 (0.00-12.00) & 61.00 \\
\hline cereal bars & 8.14 (0.00-420.00) & 0.00 (0.00-0.00 ) & 24.47 \\
\hline halva & 0.39 (0.00-140.00) & 0.00 (0.00-0.00) & 5.12 \\
\hline water ice & 12.52(0.00-1050.00) & 0.00 (0.00-0.00 ) & 18.70 \\
\hline any chocolates & 12.99(0.00-98.00) & 7.00 (3.50-21.00) & 84.75 \\
\hline chocolates bars & 31.13(0.00-700.00) & 25.00(0.00-25.00) & 51.45 \\
\hline plain chocolate & 36.45(0.00-350.00) & 12.50(12.50-75.00) & 78.92 \\
\hline vegetable oil & \(31.39(0.00-154.00)\) & \(33.00(5.50-60.50)\) & 77.71 \\
\hline sunflower oil & 13.55(0.00-154.00) & 5.50 (0.00-11.00) & 53.73 \\
\hline olive oil & \(32.50(0.00-154.00)\) & \(33.00(5.50-60.50)\) & 81.33 \\
\hline any margarine & 47.21(0.00-140.00) & 30.00(0.00-70.00) & 66.89 \\
\hline low fat margarine & 19.58(0.00-140.00) & 0.00 (0.00-10.00) & 35.58 \\
\hline half fat margarine & 15.96(0.00-140.00) & 0.00 (0.00-5.00) & 30.24 \\
\hline normal margarine & 12.36(0.00-140.00) & 0.00 (0.00-10.00) & 36.69 \\
\hline blended spreads & 18.20(0.00-140.00) & 0.00 (0.00-10.00) & 32.16 \\
\hline soya based spreads & 1.88 (0.00-140.00) & 0.00 (0.00-0.00) & 5.05 \\
\hline any butter & 29.09(0.00-140.00) & 5.00 (0.00-55.00) & 60.83 \\
\hline low fat butter & 5.34 (0.00-140.00) & 0.00 (0.00-0.00) & 14.79 \\
\hline half fat butter & 4.82 (0.00-140.00) & 0.00 (0.00-0.00) & 12.90 \\
\hline full fat butter & \(21.39(0.00-140.00)\) & 5.00 (0.00-30.00) & 50.31 \\
\hline lard & 0.09 (0.00-14.00) & 0.00 (0.00-0.00) & 6.71 \\
\hline any nuts & 14.05(0.00-182.00) & 6.50 (0.00-13.00) & 67.12 \\
\hline peanuts & 4.92 (0.00-140.00) & 0.00 (0.00-5.00) & 49.23 \\
\hline cashew nuts & 4.85 (0.00-140.00) & 0.00 (0.00-5.00) & 44.22 \\
\hline nut based spreads & 2.23 (0.00-140.00) & 0.00 (0.00-0.00) & 15.71 \\
\hline any legumes & 128.11(0.00-1680.00) & 60.00(0.00-120.00) & 70.38 \\
\hline kidney & 35.81(0.00-1260.00) & 0.00 (0.00-45.00) & 46.53 \\
\hline lentils & 28.05(0.00-1680.00) & 0.00 (0.00-60.00) & 28.64 \\
\hline chickpeas & 30.44(0.00-1260.00) & 0.00 (0.00-45.00) & 38.94 \\
\hline clusterbeans & 4.74 (0.00-630.00) & 0.00 (0.00-0.00) & 6.88 \\
\hline frenchbeans & 41.28(0.00-1260.00) & 45.00(0.00-45.00) & 52.43 \\
\hline favabeans & 16.69(0.00-1260.00) & 0.00 (0.00-0.00) & 23.26 \\
\hline soyabeans & 14.01(0.00-1260.00) & 0.00 (0.00-0.00) & 13.59 \\
\hline any leafy vegetables & 475.33(0.00-1120.00) & 440.00(240.00-560.00) & 93.29 \\
\hline lettuce & 247.03(0.00-1120.00) & 240.00(80.00-440.00) & 92.54 \\
\hline spinach & 57.26(0.00-1260.00) & 45.00(0.00-45.00) & 67.45 \\
\hline chard & 26.40(0.00-1120.00) & 0.00 (0.00-0.00) & 22.61 \\
\hline fenugreek & 1.90 (0.00-440.00) & 0.00 (0.00-0.00) & 2.44 \\
\hline wildgreens & 13.22(0.00-1260.00) & 0.00 (0.00-0.00) & 14.66 \\
\hline okra & 1.19 (0.00-330.00) & 0.00 (0.00-0.00) & 2.35 \\
\hline
\end{tabular}
\begin{tabular}{|c|c|c|c|}
\hline caper & 12.52(0.00-560.00) & 0.00 (0.00-0.00 ) & 21.60 \\
\hline tomato & 332.51(0.00-1190.00) & 255.00(85.00-467.50) & 94.75 \\
\hline aubergine & 44.91(0.00-1820.00) & 0.00 (0.00-65.00) & 39.95 \\
\hline courgette & 58.86(0.00-1400.00) & 50.00(0.00-50.00) & 55.65 \\
\hline sweet pepper & 309.04(0.00-2240.00) & 160.00(80.00-480.00) & 87.29 \\
\hline cucumber & 66.78(0.00-322.00) & 69.00(11.50-69.00) & 89.87 \\
\hline bitter mellon & 8.86 (0.00-1120.00) & 0.00 (0.00-0.00 ) & 6.06 \\
\hline carrot & 150.07(0.00-840.00) & 60.00(30.00-180.00) & 92.15 \\
\hline parsnip & 20.98(0.00-910.00) & 0.00 (0.00-32.50) & 30.27 \\
\hline turnsip & 16.31(0.00-770.00) & 0.00 (0.00-27.50) & 28.25 \\
\hline artichoke & 4.11 (0.00-275.00) & 0.00 (0.00-0.00) & 14.34 \\
\hline radish & 3.97 (0.00-112.00) & 0.00 (0.00-4.00) & 39.04 \\
\hline beetroot & 22.66(0.00-560.00) & 20.00(0.00-20.00) & 56.99 \\
\hline celery & 23.36(0.00-700.00) & 0.00 (0.00-25.00) & 42.10 \\
\hline coleslaw & 20.59(0.00-980.00) & 0.00 (0.00-35.00) & 30.34 \\
\hline sweetcorn & 56.50(0.00-1190.00) & 42.50(0.00-42.50) & 59.34 \\
\hline asparagus & 45.92(0.00-875.00) & 0.00 (0.00-62.50) & 48.39 \\
\hline herbs & 58.17(0.00-420.00) & 30.00(15.00-90.00) & 81.36 \\
\hline leek & 90.94(0.00-1260.00) & 45.00(45.00-90.00) & 79.18 \\
\hline white mushroom & 68.63(0.00-1120.00) & 40.00(40.00-80.00) & 77.78 \\
\hline onion & 215.40(0.00-980.00) & 210.00(70.00-385.00) & 93.97 \\
\hline garlic & 6.74 (0.00-42.00) & 3.00 (1.50-9.00) & 85.04 \\
\hline cauliflower & 71.01(0.00-1260.00) & 45.00(45.00-90.00) & 75.69 \\
\hline pumpkin & 20.03(0.00-1330.00) & 0.00 (0.00-0.00) & 19.26 \\
\hline brussel spouts & 29.25(0.00-1260.00) & 0.00 (0.00-45.00) & 39.95 \\
\hline broccoli & 78.99(0.00-1190.00) & 42.50(42.50-85.00) & 77.84 \\
\hline cabbage & 81.93(0.00-1330.00) & 47.50(0.00-95.00) & 68.13 \\
\hline stuffed vegetables & 12.35(0.00-1260.00) & 0.00 (0.00-0.00) & 19.81 \\
\hline pickled vegetables & 63.89(0.00-1260.00) & 45.00(0.00-45.00) & 55.78 \\
\hline ginger & \(25.05(0.00-840.00)\) & 0.00 (0.00-30.00) & 33.37 \\
\hline potatoes & 523.57(0.00-2240.00) & 480.00(160.00-880.00) & 93.68 \\
\hline mashed potato & 48.85(0.00-840.00) & 30.00(30.00-60.00) & 75.40 \\
\hline casserole & 68.76(0.00-1190.00) & 42.50(0.00-85.00) & 74.36 \\
\hline chips/fries & 89.84(0.00-2310.00) & 82.50(0.00-82.50) & 61.55 \\
\hline salads & 24.89(0.00-1190.00) & 0.00 (0.00-42.50) & 45.16 \\
\hline potato dumplings & 5.77 (0.00-560.00) & 0.00 (0.00-0.00) & 23.33 \\
\hline potato tortilla & 8.65 (0.00-1400.00) & 0.00 (0.00-0.00) & 13.95 \\
\hline sweet potato & 21.00(0.00-1120.00) & 0.00 (0.00-0.00) & 17.24 \\
\hline apple & 266.79(0.00-1050.00) & 225.00(37.50-412.50) & 88.69 \\
\hline pear & 234.89(0.00-2100.00) & 75.00(75.00-450.00) & 74.88 \\
\hline banana & 252.55(0.00-1400.00) & 100.00(50.00-300.00) & 86.41 \\
\hline peach & 105.27(0.00-1540.00) & \(55.00(0.00-110.00)\) & 58.85 \\
\hline avocado & 48.65(0.00-1260.00) & 0.00 (0.00-45.00) & 45.49 \\
\hline cherry & 51.76(0.00-1330.00) & 0.00 (0.00-47.50) & 37.54 \\
\hline
\end{tabular}
\begin{tabular}{|c|c|c|c|}
\hline rhubarb & 29.36(0.00-980.00) & 0.00 (0.00-70.00) & 26.91 \\
\hline forest fruits & 195.60(0.00-9100.00) & \(70.00(0.00-140.00)\) & 73.64 \\
\hline melon watermelon & 149.04(0.00-2800.00) & 100.00(0.00-100.00) & 64.78 \\
\hline grape & 119.30(0.00-1400.00) & 50.00(50.00-100.00) & 79.93 \\
\hline mango & 12.64(0.00-560.00) & 0.00 (0.00-20.00) & 31.54 \\
\hline apricot & 12.60(0.00-560.00) & 0.00 (0.00-20.00) & 27.11 \\
\hline nectarine & 74.96(0.00-1260.00) & 45.00(0.00-90.00) & 57.51 \\
\hline plum & 41.23(0.00-770.00) & \(27.50(0.00-27.50)\) & 50.90 \\
\hline squeezed fruit & 57.69(0.00-700.00) & 25.00(0.00-50.00) & 50.24 \\
\hline pineapple & 43.98(0.00-1120.00) & 40.00(0.00-40.00) & 55.69 \\
\hline kiwi & 44.52(0.00-840.00) & \(30.00(0.00-30.00)\) & 57.71 \\
\hline lemon & 9.50 (0.00-140.00) & 5.00 (0.00-10.00) & 57.84 \\
\hline orange & 267.32(0.00-2240.00) & 80.00(80.00-480.00) & 79.80 \\
\hline mandarine/tangarine & 169.04(0.00-1400.00) & 50.00(50.00-300.00) & 78.27 \\
\hline grapefruit & 30.59(0.00-1120.00) & 0.00 (0.00-40.00) & 29.68 \\
\hline canned fruits & 20.11(0.00-1260.00) & 0.00 (0.00-45.00) & 29.16 \\
\hline raisins & 21.46(0.00-420.00) & 0.00 (0.00-15.00) & 43.70 \\
\hline fig & 12.17(0.00-770.00) & 0.00 (0.00-0.00) & 20.95 \\
\hline prune & 13.16(0.00-840.00) & 0.00 (0.00-0.00) & 19.22 \\
\hline olive & 62.09(0.00-980.00) & \(35.00(0.00-70.00)\) & 58.36 \\
\hline juice with sugar & 130.43(0.00-2240.00) & 0.00 (0.00-80.00) & 40.14 \\
\hline juice without sugar & 207.81(0.00-2240.00) & 80.00(0.00-160.00) & 49.98 \\
\hline soft drinks & \(378.23(0.00-3500.00)\) & 0.00 (0.00-250.00) & 48.39 \\
\hline tap water & 1863.43(0.00-2800.00) & 2800.00(600.00-2800.00) & 82.83 \\
\hline mineral water & 873.34(0.00-2800.00) & 200.00(100.00-1400.00) & 75.76 \\
\hline soda with sugar & 107.23(0.00-2800.00) & 0.00 (0.00-100.00) & 26.95 \\
\hline soda without sugar & 138.96(0.00-2800.00) & 0.00 (0.00-100.00) & 25.42 \\
\hline black tea & 653.13(0.00-2660.00) & 95.00(0.00-1330.00) & 59.89 \\
\hline coffee & 1597.57(0.00-2660.00) & 1330.00(190.00-2660.00) & 79.60 \\
\hline greek coffee & 51.31(0.00-2660.00) & 0.00 (0.00-0.00) & 6.00 \\
\hline herbal tea & 416.33(0.00-2660.00) & 95.00(0.00-570.00) & 53.08 \\
\hline beer & 384.34(0.00-3990.00) & 142.50(0.00-285.00) & 61.23 \\
\hline any wine & 189.03(0.00-1750.00) & \(62.50(0.00-375.00)\) & 65.92 \\
\hline red wine & 164.81(0.00-1750.00) & \(62.50(0.00-125.00)\) & 61.81 \\
\hline white wine & 80.70(0.00-1750.00) & \(62.50(0.00-62.50)\) & 50.86 \\
\hline rose wine & 26.52(0.00-1750.00) & 0.00 (0.00-0.00) & 21.28 \\
\hline fortified wines & 8.12 (0.00-700.00) & 0.00 (0.00-0.00) & 18.44 \\
\hline spirits & 20.81(0.00-700.00) & 0.00 (0.00-25.00) & 37.21 \\
\hline any red meat & 220.74(0.00-1050.00) & \(225.00(75.00-225.00)\) & 88.33 \\
\hline roast beef steak & 123.96(0.00-1680.00) & 60.00(60.00-120.00) & 76.31 \\
\hline beef burger & 55.76(0.00-1680.00) & 60.00(0.00-60.00) & 56.83 \\
\hline minced beef meat & 156.10(0.00-1960.00) & 140.00(70.00-140.00) & 86.18 \\
\hline meat stew & 71.10(0.00-1960.00) & \(70.00(0.00-70.00)\) & 61.62 \\
\hline pork steak & 82.35(0.00-1260.00) & 45.00(45.00-90.00) & 80.42 \\
\hline
\end{tabular}
\begin{tabular}{|c|c|c|c|}
\hline meat pies & 24.36(0.00-1960.00) & 0.00 (0.00-0.00 ) & 23.10 \\
\hline sausages & 42.80(0.00-840.00) & 30.00(0.00-30.00) & 56.34 \\
\hline meat spreads & 2.76 (0.00-140.00) & 0.00 (0.00-0.00) & 12.58 \\
\hline veal & 27.57(0.00-630.00) & 0.00 (0.00-45.00) & 34.51 \\
\hline small game & 18.32(0.00-2450.00) & 0.00 (0.00-0.00 ) & 14.73 \\
\hline other game & 23.18(0.00-1225.00) & 0.00 (0.00-0.00) & 17.07 \\
\hline lamb & 19.62(0.00-630.00) & 0.00 (0.00-45.00) & 33.24 \\
\hline cured pork & 42.00(0.00-630.00) & 22.50(0.00-45.00) & 52.10 \\
\hline Salami/gammon/ham & 13.56(0.00-168.00) & 6.00 (0.00-12.00) & 71.36 \\
\hline frankfurter & 82.43(0.00-1960.00) & 70.00 (0.00-70.00) & 57.80 \\
\hline bacon cubes/bacon & 21.79(0.00-560.00) & 20.00(0.00-20.00) & 65.53 \\
\hline smoked lamb & 1.36 (0.00-330.00) & 0.00 (0.00-0.00 ) & 3.32 \\
\hline smoked game & 3.27 (0.00-980.00) & 0.00 (0.00-0.00) & 6.45 \\
\hline any poultry & 145.16(0.00-1400.00) & 100.00(50.00-300.00) & 86.25 \\
\hline chicken boiled roast & 122.34(0.00-1400.00) & 100.00(50.00-100.00) & 90.91 \\
\hline chicken in stews & 45.87(0.00-1260.00) & 45.00(0.00-45.00) & 56.08 \\
\hline turkey roasted boile & 28.95(0.00-1260.00) & 0.00 (0.00-45.00) & 34.54 \\
\hline stews breadcrumbs & 10.81(0.00-770.00) & 0.00 (0.00-0.00) & 13.36 \\
\hline smoked poultry & 14.04(0.00-980.00) & 0.00 (0.00-0.00 ) & 18.74 \\
\hline liver pates & 23.17(0.00-560.00) & 0.00 (0.00-20.00) & 45.06 \\
\hline other offal & 1.30 (0.00-137.50) & 0.00 (0.00-0.00) & 8.86 \\
\hline fresh fatty fish & 121.63(0.00-1680.00) & 60.00(60.00-120.00) & 83.84 \\
\hline fresh white fish & 81.25(0.00-1540.00) & 55.00(0.00-110.00) & 72.21 \\
\hline other fish seafood & 67.54(0.00-2240.00) & 0.00 (0.00-80.00) & 41.41 \\
\hline fresh crustaceans mollucks & 12.68(0.00-280.00) & 0.00 (0.00-20.00) & 48.19 \\
\hline smoked fatty fish & 27.51(0.00-490.00) & \(35.00(0.00-35.00)\) & 54.38 \\
\hline smoked white fish & 12.70(0.00-525.00) & 0.00 (0.00-0.00) & 19.32 \\
\hline tinned fatty fish & 29.07(0.00-980.00) & \(35.00(0.00-35.00)\) & 52.79 \\
\hline tinned crustaceans mollucks & 6.18 (0.00-275.00) & 0.00 (0.00-0.00) & 20.30 \\
\hline All eggs & 94.87(0.00-840.00) & 60.00(30.00-180.00) & 89.44 \\
\hline egg based dishes & 68.21(0.00-1540.00) & 55.00(0.00-110.00) & 68.23 \\
\hline egg based desserts & 35.84(0.00-1680.00) & 0.00 (0.00-60.00) & 35.39 \\
\hline total milk & 1117.37(0.00-2800.00) & 1100.00(100.00-1400.00) & 81.04 \\
\hline sour milk & 42.22(0.00-420.00) & 0.00 (0.00-30.00) & 46.33 \\
\hline full fat milk & 176.51(0.00-2800.00) & 0.00 (0.00-100.00) & 25.25 \\
\hline semi skimmed milk & 592.44(0.00-2800.00) & 100.00(0.00-1100.00) & 56.24 \\
\hline skimmed milk & 311.02(0.00-2800.00) & 0.00 (0.00-100.00) & 28.35 \\
\hline condensed milk & 16.96(0.00-700.00) & 0.00 (0.00-0.00 ) & 9.29 \\
\hline total yoghurt & 290.42(0.00-1750.00) & 125.00(0.00-375.00) & 70.67 \\
\hline greek style yogurt & 59.17(0.00-1750.00) & 0.00 (0.00-62.50) & 31.96 \\
\hline fromage frais & 56.28(0.00-1190.00) & 0.00 (0.00-42.50) & 42.49 \\
\hline soya milk & 22.16 (0.00-1750.00) & 0.00 (0.00-0.00) & 7.04 \\
\hline viili & 35.49(0.00-1750.00) & 0.00 (0.00-0.00) & 10.39 \\
\hline tofu & 5.02 (0.00-1120.00) & 0.00 (0.00-0.00) & 6.03 \\
\hline
\end{tabular}
\begin{tabular}{|l|l|l|r|}
\hline any cheese & \(192.25(0.00-560.00)\) & \(120.00(40.00-280.00)\) & 91.63 \\
\hline hard cheeses & \(43.11(0.00-560.00)\) & \(20.00(0.00-40.00)\) & 60.31 \\
\hline soft cheeses & \(46.03(0.00-560.00)\) & \(20.00(0.00-40.00)\) & 73.25 \\
\hline semi hard cheeses & \(130.42(0.00-560.00)\) & \(120.00(20.00-220.00)\) & 80.55 \\
\hline cottage cheese & \(25.74(0.00-560.00)\) & \(0.00(0.00-20.00)\) & 40.47 \\
\hline greek cheeses & \(3.55(0.00-560.00)\) & \(0.00(0.00-0.00)\) & 6.91 \\
\hline fresh cheeses & \(29.05(0.00-560.00)\) & \(20.00(0.00-40.00)\) & 66.05 \\
\hline ice cream & \(50.17(0.00-1050.00)\) & \(37.50(0.00-37.50)\) & 72.01 \\
\hline single cream & \(4.78(0.00-210.00)\) & \(0.00(0.00-7.50)\) & 26.49 \\
\hline creme fraiche & \(8.64(0.00-210.00)\) & \(7.50(0.00-7.50)\) & 56.47 \\
\hline sour cream & \(4.94(0.00-105.00)\) & \(0.00(0.00-7.50)\) & 40.47 \\
\hline Double clotted cream & \(5.07(0.00-210.00)\) & \(0.00(0.00-7.50)\) & 35.78 \\
\hline dressing sauces & \(12.68(0.00-280.00)\) & \(0.00(0.00-10.00)\) & 48.65 \\
\hline mayonnaise & \(9.05(0.00-210.00)\) & \(7.50(0.00-7.50)\) & 55.07 \\
\hline white sauce & \(4.44(0.00-105.00)\) & \(0.00(0.00-7.50)\) & 42.82 \\
\hline ketchup & \(12.80(0.00-210.00)\) & \(7.50(0.00-15.00)\) & 66.47 \\
\hline fresh vegetable & \(270.74(0.00-3080.00)\) & \(110.00(0.00-220.00)\) & 69.60 \\
\hline fresh meat & \(55.13(0.00-1540.00)\) & \(0.00(0.00-110.00)\) & 33.07 \\
\hline pizza & \(45.61(0.00-770.00)\) & \(55.00(0.00-55.00)\) & 67.25 \\
\hline moussaka & \(6.12(0.00-660.00)\) & \(0.00(0.00-0.00)\) & 9.51 \\
\hline
\end{tabular}

\section*{XI. Description of mean and median intake (grams per day) for each food item that was identified as significant from adjusted ESFA (FDR \(=5 \%\) ) at \(2^{\text {nd }}\) step by each country.}
\begin{tabular}{|c|c|c|c|c|}
\hline Food item & mean & median & number of individuals with non zero consumption of the specific food item & area \\
\hline any legumes & 95.75(0.00-840.00) & 60.00(60.00-120.00) & 110 & Belgium \\
\hline any legumes & 87.29(0.00-840.00) & 60.00(0.00-60.00) & 209 & Denmark \\
\hline any legumes & 72.77(0.00-840.00) & 60.00(0.00-60.00) & 84 & Finaland \\
\hline any legumes & 57.97(0.00-660.00) & 60.00(0.00-60.00) & 122 & North Germany \\
\hline any legumes & 69.28(0.00-840.00) & 60.00(0.00-60.00) & 139 & South Germany \\
\hline any legumes & 112.47(0.00-840.00) & 60.00(60.00-120.00) & 172 & Holland \\
\hline any legumes & 294.21(0.00-1680.00) & 120.00(60.00-360.00) & 215 & Portugal \\
\hline any legumes & 116.57(0.00-1680.00) & 60.00(60.00-120.00) & 183 & Poland \\
\hline any legumes & 99.65(0.00-1680.00) & 60.00(0.00-120.00) & 89 & UK \\
\hline any legumes & 144.44(0.00-1680.00) & 60.00(0.00-120.00) & 837 & Sweden \\
\hline beer & 539.74(0.00-3990.00) & 142.50(0.00-855.00) & 92 & Belgium \\
\hline beer & 516.06(0.00-3990.00) & 142.50(0.00-855.00) & 253 & Denmark \\
\hline beer & 340.16(0.00-3990.00) & 142.50(0.00-285.00) & 85 & Finaland \\
\hline beer & 402.54(0.00-3990.00) & 142.50(0.00-285.00) & 98 & North Germany \\
\hline beer & 405.46(0.00-3990.00) & 142.50(0.00-285.00) & 114 & South Germany \\
\hline beer & 503.06(0.00-3990.00) & 142.50(0.00-285.00) & 114 & Holland \\
\hline beer & 237.68(0.00-3990.00) & 0.00 (0.00-285.00) & 119 & Portugal \\
\hline beer & 308.75(0.00-3990.00) & 142.50(0.00-285.00) & 130 & Poland \\
\hline beer & 412.50(0.00-3990.00) & 0.00 (0.00-285.00) & 82 & UK \\
\hline beer & 344.98(0.00-3990.00) & 142.50(0.00-285.00) & 792 & Sweden \\
\hline bitter mellon & 1.10 (0.00-80.00) & 0.00 (0.00-0.00) & 2 & Belgium \\
\hline bitter mellon & 10.73(0.00-440.00) & 0.00 (0.00-0.00) & 37 & Denmark \\
\hline bitter mellon & 22.71(0.00-1120.00) & 0.00 (0.00-0.00) & 13 & Finaland \\
\hline bitter mellon & 4.52 (0.00-560.00) & 0.00 (0.00-0.00) & 6 & North Germany \\
\hline bitter mellon & 2.27 (0.00-240.00) & 0.00 (0.00-0.00 ) & 5 & South Germany \\
\hline bitter mellon & 3.72 (0.00-240.00) & 0.00 (0.00-0.00 ) & 14 & Holland \\
\hline bitter mellon & 34.90(0.00-1120.00) & 0.00 (0.00-0.00) & 58 & Portugal \\
\hline bitter mellon & 22.29(0.00-1120.00) & 0.00 (0.00-0.00) & 14 & Poland \\
\hline
\end{tabular}
\begin{tabular}{|c|c|c|c|c|}
\hline bitter mellon & 4.21 (0.00-240.00) & 0.00 (0.00-0.00) & 8 & UK \\
\hline bitter mellon & 2.65 (0.00-560.00) & 0.00 (0.00-0.00) & 29 & Sweden \\
\hline cabbage & 52.38(0.00-522.50) & 47.50(0.00-47.50) & 105 & Belgium \\
\hline cabbage & 57.03(0.00-665.00) & 47.50(0.00-47.50) & 236 & Denmark \\
\hline cabbage & 106.03(0.00-1330.00) & 47.50(0.00-95.00) & 109 & Finaland \\
\hline cabbage & 46.16(0.00-522.50) & 47.50(0.00-47.50) & 125 & North Germany \\
\hline cabbage & 49.21(0.00-665.00) & 47.50(0.00-47.50) & 131 & South Germany \\
\hline cabbage & 49.71(0.00-285.00) & 47.50(0.00-47.50) & 156 & Holland \\
\hline cabbage & 131.50(0.00-1330.00) & 47.50(0.00-95.00) & 153 & Portugal \\
\hline cabbage & 152.23(0.00-1330.00) & 95.00(47.50-95.00) & 199 & Poland \\
\hline cabbage & 90.83(0.00-522.50) & 47.50(47.50-95.00) & 133 & UK \\
\hline cabbage & 81.83(0.00-1330.00) & 47.50(0.00-95.00) & 744 & Sweden \\
\hline cherry & 74.83(0.00-1330.00) & 47.50(0.00-47.50) & 75 & Belgium \\
\hline cherry & 42.00(0.00-1330.00) & 0.00 (0.00-47.50) & 94 & Denmark \\
\hline cherry & 21.76(0.00-1045.00) & 0.00 (0.00-0.00) & 33 & Finaland \\
\hline cherry & 76.75(0.00-1330.00) & 47.50(0.00-95.00) & 92 & North Germany \\
\hline cherry & 64.88(0.00-1330.00) & 0.00 (0.00-47.50) & 90 & South Germany \\
\hline cherry & 39.77(0.00-1330.00) & 0.00 (0.00-47.50) & 103 & Holland \\
\hline cherry & 129.48(0.00-1330.00) & 47.50(0.00-95.00) & 157 & Portugal \\
\hline cherry & 146.57(0.00-1330.00) & 47.50(0.00-285.00) & 141 & Poland \\
\hline cherry & 41.94(0.00-665.00) & 0.00 (0.00-47.50) & 80 & UK \\
\hline cherry & 19.43(0.00-1330.00) & 0.00 (0.00-0.00) & 287 & Sweden \\
\hline condensed milk & 0.51 (0.00-25.00) & 0.00 (0.00-0.00) & 3 & Belgium \\
\hline condensed milk & 3.04 (0.00-350.00) & 0.00 (0.00-0.00) & 7 & Denmark \\
\hline condensed milk & 1.77 (0.00-150.00) & 0.00 (0.00-0.00) & 5 & Finaland \\
\hline condensed milk & 85.73(0.00-700.00) & 0.00 (0.00-25.00) & 52 & North Germany \\
\hline condensed milk & 135.31(0.00-700.00) & 0.00 (0.00-275.00) & 78 & South Germany \\
\hline condensed milk & 4.65 (0.00-350.00) & 0.00 (0.00-0.00) & 4 & Holland \\
\hline condensed milk & 5.02 (0.00-50.00) & 0.00 (0.00-0.00) & 49 & Portugal \\
\hline condensed milk & 24.88(0.00-700.00) & 0.00 (0.00-0.00) & 50 & Poland \\
\hline condensed milk & 2.78 (0.00-150.00) & 0.00 (0.00-0.00) & 12 & UK \\
\hline condensed milk & 0.85 (0.00-350.00) & 0.00 (0.00-0.00) & 25 & Sweden \\
\hline couscous & 24.66(0.00-150.00) & 0.00 (0.00-75.00) & 43 & Belgium \\
\hline couscous & 6.78 (0.00-450.00) & 0.00 (0.00-0.00) & 26 & Denmark \\
\hline couscous & 28.55(0.00-450.00) & 0.00 (0.00-75.00) & 41 & Finaland \\
\hline couscous & 21.61(0.00-150.00) & 0.00 (0.00-75.00) & 46 & North Germany \\
\hline couscous & 11.98(0.00-150.00) & 0.00 (0.00-0.00) & 29 & South Germany \\
\hline couscous & 13.26(0.00-150.00) & 0.00 (0.00-0.00 ) & 34 & Holland \\
\hline
\end{tabular}
\begin{tabular}{|c|c|c|c|c|}
\hline couscous & 29.54(0.00-1050.00) & 0.00 (0.00-0.00 ) & 32 & Portugal \\
\hline couscous & 46.43(0.00-2100.00) & 0.00 (0.00-75.00) & 67 & Poland \\
\hline couscous & 25.88(0.00-450.00) & 0.00 (0.00-75.00) & 43 & UK \\
\hline couscous & 32.33(0.00-450.00) & 0.00 (0.00-75.00) & 392 & Sweden \\
\hline crisp fried Cakes & 0.68 (0.00-60.00) & 0.00 (0.00-0.00) & 3 & Belgium \\
\hline crisp fried Cakes & 0.45 (0.00-140.00) & 0.00 (0.00-0.00) & 3 & Denmark \\
\hline crisp fried Cakes & 0.00 (0.00-0.00) & 0.00 (0.00-0.00) & 0 & Finaland \\
\hline crisp fried Cakes & 1.13 (0.00-60.00) & 0.00 (0.00-0.00) & 14 & North Germany \\
\hline crisp fried Cakes & 0.57 (0.00-20.00) & 0.00 (0.00-0.00) & 10 & South Germany \\
\hline crisp fried Cakes & 0.05 (0.00-10.00) & 0.00 (0.00-0.00) & 1 & Holland \\
\hline crisp fried Cakes & 1.24 (0.00-20.00) & 0.00 (0.00-0.00) & 27 & Portugal \\
\hline crisp fried Cakes & 2.00 (0.00-110.00) & 0.00 (0.00-0.00) & 27 & Poland \\
\hline crisp fried Cakes & 0.53 (0.00-60.00) & 0.00 (0.00-0.00) & 3 & UK \\
\hline crisp fried Cakes & 0.08 (0.00-20.00) & 0.00 (0.00-0.00) & 7 & Sweden \\
\hline greek style yogurt & 56.93(0.00-1750.00) & 0.00 (0.00-0.00) & 29 & Belgium \\
\hline greek style yogurt & 34.07(0.00-875.00) & 0.00 (0.00-62.50) & 112 & Denmark \\
\hline greek style yogurt & 37.50(0.00-875.00) & 0.00 (0.00-0.00) & 33 & Finaland \\
\hline greek style yogurt & 66.74(0.00-875.00) & 0.00 (0.00-62.50) & 45 & North Germany \\
\hline greek style yogurt & 61.21(0.00-1750.00) & 0.00 (0.00-0.00 ) & 44 & South Germany \\
\hline greek style yogurt & 53.49(0.00-875.00) & 0.00 (0.00-62.50) & 56 & Holland \\
\hline greek style yogurt & 40.06(0.00-1750.00) & 0.00 (0.00-0.00) & 25 & Portugal \\
\hline greek style yogurt & 71.73(0.00-1750.00) & 0.00 (0.00-62.50) & 70 & Poland \\
\hline greek style yogurt & 73.46(0.00-875.00) & 0.00 (0.00-62.50) & 71 & UK \\
\hline greek style yogurt & 69.30(0.00-1750.00) & 0.00 (0.00-62.50) & 496 & Sweden \\
\hline kidney & 11.71(0.00-90.00) & 0.00 (0.00-0.00) & 31 & Belgium \\
\hline kidney & 10.55(0.00-270.00) & 0.00 (0.00-0.00) & 73 & Denmark \\
\hline kidney & 22.94(0.00-270.00) & 0.00 (0.00-45.00) & 58 & Finaland \\
\hline kidney & 14.24(0.00-90.00) & 0.00 (0.00-45.00) & 52 & North Germany \\
\hline kidney & 25.52(0.00-270.00) & 0.00 (0.00-45.00) & 85 & South Germany \\
\hline kidney & 29.72(0.00-270.00) & 45.00(0.00-45.00) & 114 & Holland \\
\hline kidney & 81.49(0.00-630.00) & 45.00(45.00-90.00) & 209 & Portugal \\
\hline kidney & 45.64(0.00-630.00) & 45.00(0.00-45.00) & 132 & Poland \\
\hline kidney & 32.11(0.00-270.00) & 45.00(0.00-45.00) & 86 & UK \\
\hline kidney & 42.90(0.00-1260.00) & 22.50(0.00-45.00) & 588 & Sweden \\
\hline lemon & 7.36 (0.00-70.00) & 0.00 (0.00-5.00) & 69 & Belgium \\
\hline lemon & 7.58 (0.00-70.00) & 5.00 (0.00-5.00) & 188 & Denmark \\
\hline lemon & 5.26 (0.00-140.00) & 0.00 (0.00-5.00) & 72 & Finaland \\
\hline lemon & 8.95 (0.00-70.00) & 5.00 (0.00-10.00) & 101 & North Germany \\
\hline lemon & 5.72 (0.00-140.00) & 0.00 (0.00-5.00) & 90 & South \\
\hline
\end{tabular}

Appendices
\begin{tabular}{|c|c|c|c|c|}
\hline & & & & Germany \\
\hline lemon & 5.28 (0.00-70.00) & 0.00 (0.00-5.00 ) & 96 & Holland \\
\hline lemon & 8.98 (0.00-140.00) & 0.00 (0.00-10.00) & 128 & Portugal \\
\hline lemon & 35.88(0.00-140.00) & 10.00(5.00-55.00) & 181 & Poland \\
\hline lemon & 7.40 (0.00-70.00) & 5.00 (0.00-10.00) & 89 & UK \\
\hline lemon & 8.09 (0.00-140.00) & 5.00 (0.00-10.00) & 761 & Sweden \\
\hline lentils & 18.08(0.00-120.00) & 0.00 (0.00-0.00) & 35 & Belgium \\
\hline lentils & 6.10 (0.00-360.00) & 0.00 (0.00-0.00) & 24 & Denmark \\
\hline lentils & 11.61(0.00-120.00) & 0.00 (0.00-0.00) & 27 & Finaland \\
\hline lentils & 34.92(0.00-120.00) & 60.00(0.00-60.00) & 96 & North Germany \\
\hline lentils & 35.88(0.00-360.00) & 0.00 (0.00-60.00) & 96 & South Germany \\
\hline lentils & 9.49 (0.00-360.00) & 0.00 (0.00-0.00) & 28 & Holland \\
\hline lentils & 13.20(0.00-360.00) & 0.00 (0.00-0.00) & 40 & Portugal \\
\hline lentils & 10.00(0.00-360.00) & 0.00 (0.00-0.00) & 18 & Poland \\
\hline lentils & 40.70(0.00-360.00) & 0.00 (0.00-60.00) & 59 & UK \\
\hline lentils & 43.78(0.00-1680.00) & 0.00 (0.00-60.00) & 456 & Sweden \\
\hline mango & 10.41(0.00-120.00) & 0.00 (0.00-20.00) & 51 & Belgium \\
\hline mango & 5.37 (0.00-120.00) & 0.00 (0.00-0.00) & 69 & Denmark \\
\hline mango & 3.61 (0.00-40.00) & 0.00 (0.00-0.00) & 27 & Finaland \\
\hline mango & 17.74(0.00-280.00) & 0.00 (0.00-20.00) & 70 & North Germany \\
\hline mango & 8.25 (0.00-280.00) & 0.00 (0.00-0.00 ) & 44 & South Germany \\
\hline mango & 11.16(0.00-220.00) & 0.00 (0.00-20.00) & 83 & Holland \\
\hline mango & 34.05(0.00-560.00) & 20.00(0.00-40.00) & 132 & Portugal \\
\hline mango & 14.57(0.00-560.00) & 0.00 (0.00-0.00 ) & 43 & Poland \\
\hline mango & 19.42(0.00-560.00) & 0.00 (0.00-20.00) & 75 & UK \\
\hline mango & 10.48(0.00-560.00) & 0.00 (0.00-20.00) & 374 & Sweden \\
\hline moussaka & 10.27(0.00-120.00) & 0.00 (0.00-0.00) & 24 & Belgium \\
\hline moussaka & 6.27 (0.00-60.00) & 0.00 (0.00-0.00) & 37 & Denmark \\
\hline moussaka & 7.74 (0.00-60.00) & 0.00 (0.00-0.00) & 20 & Finaland \\
\hline moussaka & 2.03 (0.00-60.00) & 0.00 (0.00-0.00) & 6 & North Germany \\
\hline moussaka & 3.09 (0.00-120.00) & 0.00 (0.00-0.00 ) & 9 & South Germany \\
\hline moussaka & 5.58 (0.00-120.00) & 0.00 (0.00-0.00) & 19 & Holland \\
\hline moussaka & 4.40 (0.00-660.00) & 0.00 (0.00-0.00) & 8 & Portugal \\
\hline moussaka & 1.43 (0.00-60.00) & 0.00 (0.00-0.00) & 5 & Poland \\
\hline moussaka & 8.42 (0.00-120.00) & 0.00 (0.00-0.00) & 22 & UK \\
\hline moussaka & 7.45 (0.00-120.00) & 0.00 (0.00-0.00) & 142 & Sweden \\
\hline okra & 1.23 (0.00-30.00) & 0.00 (0.00-0.00) & 6 & Belgium \\
\hline okra & 0.17 (0.00-30.00) & 0.00 (0.00-0.00) & 2 & Denmark \\
\hline okra & 0.00 (0.00-0.00) & 0.00 (0.00-0.00) & 0 & Finaland \\
\hline okra & 0.00 (0.00-0.00) & 0.00 (0.00-0.00) & 0 & North \\
\hline
\end{tabular}
\begin{tabular}{|c|c|c|c|c|}
\hline & & & & Germany \\
\hline okra & 0.62 (0.00-30.00) & 0.00 (0.00-0.00) & 4 & South Germany \\
\hline okra & 1.12 (0.00-30.00) & 0.00 (0.00-0.00) & 8 & Holland \\
\hline okra & 3.71 (0.00-330.00) & 0.00 (0.00-0.00) & 14 & Portugal \\
\hline okra & 0.43 (0.00-60.00) & 0.00 (0.00-0.00) & 2 & Poland \\
\hline okra & 1.93 (0.00-60.00) & 0.00 (0.00-0.00) & 10 & UK \\
\hline okra & 1.40 (0.00-330.00) & 0.00 (0.00-0.00) & 26 & Sweden \\
\hline other game & 10.19(0.00-175.00) & 0.00 (0.00-0.00) & 16 & Belgium \\
\hline other game & 6.43 (0.00-175.00) & 0.00 (0.00-0.00) & 25 & Denmark \\
\hline other game & 27.10(0.00-525.00) & 0.00 (0.00-0.00) & 32 & Finaland \\
\hline other game & 10.88(0.00-175.00) & 0.00 (0.00-0.00 ) & 20 & North Germany \\
\hline other game & 7.67 (0.00-87.50) & 0.00 (0.00-0.00 ) & 17 & South Germany \\
\hline other game & 9.36 (0.00-87.50) & 0.00 (0.00-0.00 ) & 23 & Holland \\
\hline other game & 5.74 (0.00-525.00) & 0.00 (0.00-0.00) & 10 & Portugal \\
\hline other game & 3.75 (0.00-175.00) & 0.00 (0.00-0.00) & 7 & Poland \\
\hline other game & 2.05 (0.00-87.50) & 0.00 (0.00-0.00) & 4 & UK \\
\hline other game & 46.65(0.00-1225.00) & 0.00 (0.00-87.50) & 370 & Sweden \\
\hline peach & 105.48(0.00-770.00) & 55.00(0.00-110.00) & 97 & Belgium \\
\hline peach & 60.13(0.00-770.00) & 55.00(0.00-55.00) & 182 & Denmark \\
\hline peach & 42.94(0.00-605.00) & 0.00 (0.00-55.00) & 74 & Finaland \\
\hline peach & 87.32(0.00-770.00) & 55.00(0.00-110.00) & 104 & North Germany \\
\hline peach & 119.92(0.00-1540.00) & 55.00(0.00-110.00) & 134 & South Germany \\
\hline peach & 59.09(0.00-1540.00) & 55.00(0.00-55.00) & 113 & Holland \\
\hline peach & 337.64(0.00-1540.00) & 110.00(55.00-605.00) & 228 & Portugal \\
\hline peach & 144.83(0.00-1540.00) & 55.00(55.00-110.00) & 163 & Poland \\
\hline peach & 98.42(0.00-770.00) & 55.00(0.00-110.00) & 108 & UK \\
\hline peach & 78.52(0.00-1540.00) & 55.00(0.00-55.00) & 603 & Sweden \\
\hline peanuts & 3.08 (0.00-30.00) & 0.00 (0.00-5.00 ) & 57 & Belgium \\
\hline peanuts & 3.21 (0.00-70.00) & 0.00 (0.00-5.00) & 122 & Denmark \\
\hline peanuts & 3.35 (0.00-70.00) & 0.00 (0.00-5.00) & 54 & Finaland \\
\hline peanuts & 2.85 (0.00-30.00) & 0.00 (0.00-5.00) & 74 & North Germany \\
\hline peanuts & 5.28 (0.00-140.00) & 5.00 (0.00-5.00) & 100 & South Germany \\
\hline peanuts & 8.49 (0.00-70.00) & 5.00 (0.00-10.00) & 140 & Holland \\
\hline peanuts & 4.54 (0.00-55.00) & 5.00 (0.00-5.00) & 133 & Portugal \\
\hline peanuts & 4.74 (0.00-140.00) & 5.00 (0.00-5.00 ) & 106 & Poland \\
\hline peanuts & 6.35 (0.00-70.00) & 5.00 (0.00-5.00) & 86 & UK \\
\hline peanuts & 5.37 (0.00-140.00) & 5.00 (0.00-5.00 ) & 639 & Sweden \\
\hline pumpkin & 43.60(0.00-665.00) & 47.50(0.00-47.50) & 85 & Belgium \\
\hline pumpkin & 9.80 (0.00-285.00) & 0.00 (0.00-0.00 ) & 47 & Denmark \\
\hline
\end{tabular}
\begin{tabular}{|c|c|c|c|c|}
\hline pumpkin & 21.45(0.00-522.50) & 0.00 (0.00-0.00 ) & 37 & Finaland \\
\hline pumpkin & 16.10(0.00-285.00) & 0.00 (0.00-0.00) & 43 & North Germany \\
\hline pumpkin & 9.79 (0.00-522.50) & 0.00 (0.00-0.00) & 25 & South Germany \\
\hline pumpkin & 14.80(0.00-285.00) & 0.00 (0.00-0.00) & 50 & Holland \\
\hline pumpkin & 107.47(0.00-1330.00) & 47.50(0.00-95.00) & 156 & Portugal \\
\hline pumpkin & 26.01(0.00-1330.00) & 0.00 (0.00-0.00) & 47 & Poland \\
\hline pumpkin & 10.83(0.00-285.00) & 0.00 (0.00-0.00) & 29 & UK \\
\hline pumpkin & 4.24 (0.00-665.00) & 0.00 (0.00-0.00) & 72 & Sweden \\
\hline rhubarb & 28.77(0.00-980.00) & 0.00 (0.00-70.00) & 38 & Belgium \\
\hline rhubarb & 27.88(0.00-420.00) & 0.00 (0.00-70.00) & 103 & Denmark \\
\hline rhubarb & 33.87(0.00-770.00) & 0.00 (0.00-70.00) & 51 & Finaland \\
\hline rhubarb & 41.92(0.00-980.00) & 0.00 (0.00-70.00) & 53 & North Germany \\
\hline rhubarb & 39.33(0.00-980.00) & 0.00 (0.00-0.00 ) & 41 & South Germany \\
\hline rhubarb & 20.84(0.00-420.00) & 0.00 (0.00-0.00) & 52 & Holland \\
\hline rhubarb & 1.89 (0.00-140.00) & 0.00 (0.00-0.00) & 5 & Portugal \\
\hline rhubarb & 32.67(0.00-980.00) & 0.00 (0.00-0.00) & 44 & Poland \\
\hline rhubarb & 35.61(0.00-770.00) & 0.00 (0.00-70.00) & 64 & UK \\
\hline rhubarb & 31.85(0.00-980.00) & 0.00 (0.00-70.00) & 375 & Sweden \\
\hline roast beef steak & 156.99(0.00-840.00) & 60.00(0.00-120.00) & 109 & Belgium \\
\hline roast beef steak & 93.73(0.00-840.00) & 60.00(60.00-120.00) & 289 & Denmark \\
\hline roast beef steak & 110.32(0.00-360.00) & 60.00(0.00-120.00) & 114 & Finaland \\
\hline roast beef steak & 44.75(0.00-360.00) & 0.00 (0.00-60.00) & 87 & North Germany \\
\hline roast beef steak & 71.13(0.00-840.00) & 60.00(0.00-60.00) & 123 & South Germany \\
\hline roast beef steak & 237.77(0.00-840.00) & 120.00(60.00-360.00) & 197 & Holland \\
\hline roast beef steak & 259.00(0.00-1680.00) & 120.00(60.00-360.00) & 221 & Portugal \\
\hline roast beef steak & 158.86(0.00-1680.00) & 60.00(60.00-120.00) & 176 & Poland \\
\hline roast beef steak & 121.40(0.00-840.00) & 60.00(60.00-120.00) & 134 & UK \\
\hline roast beef steak & 95.00(0.00-840.00) & 60.00(60.00-120.00) & 892 & Sweden \\
\hline single cream & 7.14 (0.00-105.00) & 7.50 (0.00-7.50) & 80 & Belgium \\
\hline single cream & 6.86 (0.00-210.00) & 0.00 (0.00-0.00) & 40 & Denmark \\
\hline single cream & 2.03 (0.00-45.00) & 0.00 (0.00-0.00) & 27 & Finaland \\
\hline single cream & 5.81 (0.00-82.50) & 7.50 (0.00-7.50) & 92 & North Germany \\
\hline single cream & 6.57 (0.00-45.00) & 7.50 (0.00-7.50) & 116 & South Germany \\
\hline single cream & 6.49 (0.00-210.00) & 0.00 (0.00-7.50) & 99 & Holland \\
\hline single cream & 2.58 (0.00-45.00) & 0.00 (0.00-7.50) & 70 & Portugal \\
\hline single cream & 16.07(0.00-210.00) & 7.50 (0.00-15.00) & 142 & Poland \\
\hline single cream & 3.11 (0.00-45.00) & 0.00 (0.00-7.50) & 49 & UK \\
\hline single cream & 2.17 (0.00-210.00) & 0.00 (0.00-0.00 ) & 98 & Sweden \\
\hline
\end{tabular}

Appendices
\begin{tabular}{|c|c|c|c|c|}
\hline smoked fatty fish & 30.21(0.00-210.00) & 35.00(0.00-35.00) & 102 & Belgium \\
\hline smoked fatty fish & 23.53(0.00-385.00) & 0.00 (0.00-35.00) & 160 & Denmark \\
\hline smoked fatty fish & 30.94(0.00-385.00) & 35.00(0.00-35.00) & 94 & Finaland \\
\hline smoked fatty fish & 16.41(0.00-70.00) & 0.00 (0.00-35.00) & 72 & North Germany \\
\hline smoked fatty fish & 17.32(0.00-490.00) & 0.00 (0.00-35.00) & 76 & South Germany \\
\hline smoked fatty fish & 25.56(0.00-210.00) & 35.00(0.00-35.00) & 123 & Holland \\
\hline smoked fatty fish & 25.81(0.00-385.00) & 0.00 (0.00-35.00) & 105 & Portugal \\
\hline smoked fatty fish & 32.67(0.00-385.00) & 35.00(0.00-35.00) & 130 & Poland \\
\hline smoked fatty fish & 20.06(0.00-210.00) & 0.00 (0.00-35.00) & 74 & UK \\
\hline smoked fatty fish & 32.17(0.00-490.00) & 35.00(0.00-35.00) & 733 & Sweden \\
\hline smoked poultry & 21.34(0.00-490.00) & 0.00 (0.00-35.00) & 47 & Belgium \\
\hline smoked poultry & 4.94 (0.00-490.00) & 0.00 (0.00-0.00 ) & 34 & Denmark \\
\hline smoked poultry & 15.13(0.00-980.00) & 0.00 (0.00-0.00) & 16 & Finaland \\
\hline smoked poultry & 3.36 (0.00-35.00) & 0.00 (0.00-0.00 ) & 17 & North Germany \\
\hline smoked poultry & 7.58 (0.00-210.00) & 0.00 (0.00-0.00 ) & 24 & South Germany \\
\hline smoked poultry & 3.09 (0.00-70.00) & 0.00 (0.00-0.00 ) & 16 & Holland \\
\hline smoked poultry & 9.19 (0.00-210.00) & 0.00 (0.00-0.00) & 35 & Portugal \\
\hline smoked poultry & 12.33(0.00-210.00) & 0.00 (0.00-35.00) & 55 & Poland \\
\hline smoked poultry & 4.09 (0.00-210.00) & 0.00 (0.00-0.00 ) & 13 & UK \\
\hline smoked poultry & 23.21(0.00-980.00) & 0.00 (0.00-35.00) & 318 & Sweden \\
\hline sour cream & 0.87 (0.00-15.00) & 0.00 (0.00-0.00) & 15 & Belgium \\
\hline sour cream & 1.74 (0.00-45.00) & 0.00 (0.00-0.00) & 63 & Denmark \\
\hline sour cream & 3.39 (0.00-45.00) & 0.00 (0.00-7.50) & 62 & Finaland \\
\hline sour cream & 4.66 (0.00-82.50) & 0.00 (0.00-7.50) & 70 & North Germany \\
\hline sour cream & 4.21 (0.00-105.00) & 0.00 (0.00-7.50) & 71 & South Germany \\
\hline sour cream & 4.01 (0.00-45.00) & 0.00 (0.00-7.50) & 80 & Holland \\
\hline sour cream & 0.14 (0.00-7.50) & 0.00 (0.00-0.00) & 5 & Portugal \\
\hline sour cream & 11.64(0.00-105.00) & 7.50 (0.00-15.00) & 116 & Poland \\
\hline sour cream & 1.40 (0.00-15.00) & 0.00 (0.00-0.00 ) & 30 & UK \\
\hline sour cream & 7.32 (0.00-105.00) & 7.50 (0.00-7.50) & 730 & Sweden \\
\hline thin biscuits & 2.18 (0.00-28.00) & 0.00 (0.00-2.00) & 50 & Belgium \\
\hline thin biscuits & 2.14 (0.00-28.00) & 0.00 (0.00-2.00 ) & 152 & Denmark \\
\hline thin biscuits & 3.94 (0.00-28.00) & 2.00 (0.00-4.00) & 78 & Finaland \\
\hline thin biscuits & 2.10 (0.00-28.00) & 0.00 (0.00-2.00 ) & 81 & North Germany \\
\hline thin biscuits & 2.08 (0.00-28.00) & 0.00 (0.00-2.00 ) & 90 & \begin{tabular}{l}
South \\
Germany
\end{tabular} \\
\hline thin biscuits & 5.17 (0.00-28.00) & 2.00 (0.00-4.00 ) & 137 & Holland \\
\hline thin biscuits & 2.41 (0.00-56.00) & 0.00 (0.00-2.00 ) & 93 & Portugal \\
\hline thin biscuits & 2.00 (0.00-56.00) & 0.00 (0.00-2.00 ) & 99 & Poland \\
\hline
\end{tabular}

Appendices
\begin{tabular}{|c|c|c|c|c|}
\hline thin biscuits & 3.30 (0.00-28.00) & 2.00 (0.00-4.00 ) & 97 & UK \\
\hline thin biscuits & 8.37 (0.00-56.00) & 2.00 (2.00-12.00) & 898 & Sweden \\
\hline tofu & 10.68(0.00-560.00) & 0.00 (0.00-0.00) & 9 & Belgium \\
\hline tofu & 1.24 (0.00-240.00) & 0.00 (0.00-0.00) & 5 & Denmark \\
\hline tofu & 16.77(0.00-1120.00) & 0.00 (0.00-0.00) & 18 & Finaland \\
\hline tofu & 2.03 (0.00-80.00) & 0.00 (0.00-0.00) & 7 & North Germany \\
\hline tofu & 2.27 (0.00-240.00) & 0.00 (0.00-0.00 ) & 5 & South Germany \\
\hline tofu & 8.56 (0.00-240.00) & 0.00 (0.00-0.00) & 27 & Holland \\
\hline tofu & 4.32 (0.00-560.00) & 0.00 (0.00-0.00) & 12 & Portugal \\
\hline tofu & 3.24 (0.00-240.00) & 0.00 (0.00-0.00) & 10 & Poland \\
\hline tofu & 4.44 (0.00-240.00) & 0.00 (0.00-0.00) & 13 & UK \\
\hline tofu & 4.73 (0.00-1120.00) & 0.00 (0.00-0.00) & 79 & Sweden \\
\hline turnip & 9.04 (0.00-165.00) & 0.00 (0.00-27.50) & 38 & Belgium \\
\hline turnip & 4.97 (0.00-385.00) & 0.00 (0.00-0.00) & 36 & Denmark \\
\hline turnip & 7.10 (0.00-55.00) & 0.00 (0.00-0.00) & 38 & Finaland \\
\hline turnip & 5.59 (0.00-165.00) & 0.00 (0.00-0.00) & 31 & North Germany \\
\hline turnip & 2.84 (0.00-55.00) & 0.00 (0.00-0.00) & 18 & South Germany \\
\hline turnip & 5.50 (0.00-302.50) & 0.00 (0.00-0.00) & 31 & Holland \\
\hline turnip & 66.47(0.00-770.00) & 27.50(0.00-55.00) & 151 & Portugal \\
\hline turnip & 9.43 (0.00-770.00) & 0.00 (0.00-0.00 ) & 29 & Poland \\
\hline turnip & 40.69(0.00-302.50) & 27.50(0.00-55.00) & 107 & UK \\
\hline turnip & 14.29(0.00-385.00) & 0.00 (0.00-27.50) & 388 & Sweden \\
\hline vegetable oil & 25.84(0.00-154.00) & 5.50 (0.00-33.00) & 93 & Belgium \\
\hline vegetable oil & 31.26(0.00-154.00) & 33.00(5.50-60.50) & 282 & Denmark \\
\hline vegetable oil & 21.25(0.00-77.00) & 5.50 (0.00-33.00) & 104 & Finaland \\
\hline vegetable oil & 32.44(0.00-154.00) & 33.00(5.50-60.50) & 149 & North Germany \\
\hline vegetable oil & 42.72(0.00-154.00) & 33.00(11.00-60.50) & 179 & South Germany \\
\hline vegetable oil & 31.13(0.00-154.00) & 11.00(0.00-60.50) & 153 & Holland \\
\hline vegetable oil & 15.40(0.00-154.00) & 5.50 (0.00-11.00) & 149 & Portugal \\
\hline vegetable oil & 30.75(0.00-154.00) & 33.00(5.50-60.50) & 176 & Poland \\
\hline vegetable oil & 16.79(0.00-154.00) & 5.50 (0.00-33.00) & 99 & UK \\
\hline vegetable oil & 37.24(0.00-154.00) & 33.00(5.50-60.50) & 1001 & Sweden \\
\hline viili like fermented yoghurt & 0.86 (0.00-125.00) & 0.00 (0.00-0.00 ) & 1 & Belgium \\
\hline viili like fermented yoghurt & 3.88 (0.00-875.00) & 0.00 (0.00-0.00) & 8 & Denmark \\
\hline viili like fermented yoghurt & 89.11(0.00-875.00) & 0.00 (0.00-62.50) & 67 & Finaland \\
\hline viili like fermented yoghurt & 9.53 (0.00-687.50) & 0.00 (0.00-0.00) & 15 & North Germany \\
\hline viili like fermented & 18.36(0.00-875.00) & 0.00 (0.00-0.00) & 29 & South \\
\hline
\end{tabular}
\begin{tabular}{|l|l|l|r|l|}
\hline yoghurt & & & Germany \\
\hline \begin{tabular}{l} 
viili like fermented \\
yoghurt
\end{tabular} & \(18.90(0.00-1750.00)\) & \(0.00(0.00-0.00)\) & 9 & Holland \\
\hline \begin{tabular}{l} 
viili like fermented \\
yoghurt
\end{tabular} & \(285.23(0.00-1750.00)\) & \(62.50(0.00-375.00)\) & 24 & Poland \\
\hline \begin{tabular}{l} 
viili like fermented \\
yoghurt
\end{tabular} & \(33.93(0.00-1750.00)\) & \(0.00(0.00-0.00)\) & 3 & UK \\
\hline \begin{tabular}{l} 
viili like fermented \\
yoghurt
\end{tabular} & \(7.68(0.00-875.00)\) & \(0.00(0.00-0.00)\) & 16 & Sweden \\
\hline \begin{tabular}{l} 
viili like fermented \\
yoghurt
\end{tabular} & \(1.33(0.00-375.00)\) & \(0.00(0.00-0.00)\) & & \\
\hline
\end{tabular}

\section*{XII. Associations between dietary patterns and respiratory outcomes by smoking status, body mass index and gender: results of meta-analyses. OR; odds ratio.}
\begin{tabular}{|c|c|c|c|c|}
\hline respiratory outcomes & Dietary pattenr & odds ratio (95\%CI) & isquare ( \(p\)-value for heterogeneity) & Strata \\
\hline asthma & fruit, fish and vegetables & 1.12(0.92,1.35) & 43.84(0.066) & never smoker \\
\hline asthma & meat, potatoes and sweets & 0.99(0.88,1.11) & 0.00(0.758) & never smoker \\
\hline chronic sinusitis & fruit, fish and vegetables & 1.07(0.89,1.27) & 6.47(0.382) & never smoker \\
\hline chronic sinusitis & meat, potatoes and sweets & 1.03(0.85,1.25) & 24.68(0.216) & never smoker \\
\hline allergic rhinitis & fruit, fish and vegetables & 0.99(0.79,1.25) & 63.24(0.004) & never smoker \\
\hline allergic rhinitis & meat, potatoes and sweets & 0.99(0.85,1.14) & 25.21(0.211) & never smoker \\
\hline eczema & fruit, fish and vegetables & 1.13(0.92,1.39) & 57.10(0.013) & never smoker \\
\hline eczema & meat, potatoes and sweets & 0.93(0.78,1.11) & 40.97(0.084) & never smoker \\
\hline asthma & fruit, fish and vegetables & 1.15(0.95,1.38) & 11.64(0.336) & ever smoker \\
\hline asthma & meat, potatoes and sweets & 0.88(0.71,1.07) & 15.74(0.298) & ever smoker \\
\hline chronic sinusitis & fruit, fish and vegetables & 1.19(0.92,1.54) & 37.74(0.107) & ever smoker \\
\hline chronic sinusitis & meat, potatoes and sweets & 1.21(0.93,1.57) & 36.15(0.119) & ever smoker \\
\hline allergic rhinitis & fruit, fish and vegetables & 0.96(0.76,1.22) & 36.62(0.115) & ever smoker \\
\hline allergic rhinitis & meat, potatoes and sweets & 1.28(1.07,1.55) & 15.31(0.302) & ever smoker \\
\hline eczema & fruit, fish and vegetables & 1.35(0.99,1.83) & 64.16(0.003) & ever smoker \\
\hline eczema & meat, potatoes and sweets & 1.07(0.83,1.38) & 55.56(0.016) & ever smoker \\
\hline asthma & fruit, fish and vegetables & 0.93(0.56,1.55) & 73.33(0.000) & current smoker \\
\hline asthma & meat, potatoes and sweets & 1.00(0.67,1.47) & 46.01(0.054) & current smoker \\
\hline chronic sinusitis & fruit, fish and vegetables & 0.85(0.66,1.08) & 10.12(0.351) & current smoker \\
\hline chronic sinusitis & meat, potatoes and sweets & 1.40(0.95,2.06) & 44.31(0.064) & current smoker \\
\hline
\end{tabular}
\begin{tabular}{|c|c|c|c|c|}
\hline allergic rhinitis & fruit, fish and vegetables & 1.12(0.80,1.55) & 59.87(0.008) & current smoker \\
\hline allergic rhinitis & meat, potatoes and sweets & 1.20(0.96,1.51) & 0.00(0.536) & current smoker \\
\hline eczema & fruit, fish and vegetables & 1.01(0.73,1.38) & 40.60(0.087) & current smoker \\
\hline eczema & meat, potatoes and sweets & 1.31(1.04,1.65) & 0.00(0.852) & current smoker \\
\hline asthma & fruit, fish and vegetables & 1.10(0.94,1.27) & 7.85(0.370) & female \\
\hline asthma & meat, potatoes and sweets & 1.02(0.87,1.19) & 0.00(0.736) & female \\
\hline chronic sinusitis & fruit, fish and vegetables & 0.97(0.76,1.25) & 43.69(0.067) & female \\
\hline chronic sinusitis & meat, potatoes and sweets & 1.13(0.88,1.45) & 48.51(0.042) & female \\
\hline allergic rhinitis & fruit, fish and vegetables & 1.11(0.84,1.47) & 65.22(0.002) & female \\
\hline allergic rhinitis & meat, potatoes and sweets & 1.00(0.87,1.15) & 0.00(0.471) & female \\
\hline eczema & fruit, fish and vegetables & 1.06(0.85,1.33) & 54.14(0.020) & female \\
\hline eczema & meat, potatoes and sweets & 0.89(0.73,1.08) & 42.20(0.076) & female \\
\hline asthma & fruit, fish and vegetables & 1.10(0.95,1.26) & 22.41(0.237) & male \\
\hline asthma & meat, potatoes and sweets & 0.90(0.81,1.01) & 0.00(0.568) & male \\
\hline chronic sinusitis & fruit, fish and vegetables & 1.08(0.95,1.23) & 0.00(0.575) & male \\
\hline chronic sinusitis & meat, potatoes and sweets & 1.08(0.95,1.22) & 0.00(0.468) & male \\
\hline allergic rhinitis & fruit, fish and vegetables & 0.87(0.69,1.08) & 68.01(0.001) & male \\
\hline allergic rhinitis & meat, potatoes and sweets & 1.10(0.98,1.24) & 8.37(0.365) & male \\
\hline eczema & fruit, fish and vegetables & 1.17(0.97,1.42) & 59.37(0.008) & male \\
\hline eczema & meat, potatoes and sweets & 1.09(0.93,1.27) & 41.14(0.083) & male \\
\hline asthma & fruit, fish and vegetables & 1.12(0.97,1.31) & 5.87(0.387) & non-obese \\
\hline asthma & meat, potatoes and sweets & 0.90(0.75,1.08) & 26.80(0.197) & non-obese \\
\hline chronic sinusitis & fruit, fish and vegetables & 1.13(0.97,1.32) & 6.58(0.381) & non-obese \\
\hline chronic sinusitis & meat, potatoes and sweets & 1.00(0.80,1.25) & 48.02(0.044) & non-obese \\
\hline allergic rhinitis & fruit, fish and vegetables & 0.92(0.67,1.26) & 75.87(0.000) & non-obese \\
\hline
\end{tabular}
\begin{tabular}{|l|l|l|l|l|}
\hline allergic rhinitis & \begin{tabular}{l} 
meat, potatoes \\
and sweets
\end{tabular} & \(1.04(0.79,1.38)\) & \(68.90(0.001)\) & non-obese \\
\hline eczema & \begin{tabular}{l} 
fruit, fish and \\
vegetables
\end{tabular} & \(1.24(1.00,1.55)\) & \(45.12(0.059)\) & non-obese \\
\hline eczema & \begin{tabular}{l} 
meat, potatoes \\
and sweets
\end{tabular} & \(0.94(0.75,1.18)\) & \(56.51(0.014)\) & non-obese \\
\hline asthma & \begin{tabular}{l} 
fruit, fish and \\
vegetables
\end{tabular} & \(1.05(0.94,1.17)\) & \(0.00(0.812)\) & obese \\
\hline asthma & \begin{tabular}{l} 
meat, potatoes \\
and sweets
\end{tabular} & \(1.05(0.92,1.19)\) & \(0.00(0.771)\) & obese \\
\hline chronic sinusitis & \begin{tabular}{l} 
fruit, fish and \\
vegetables
\end{tabular} & \(1.04(0.82,1.31)\) & \(50.90(0.032)\) & obese \\
\hline chronic sinusitis & \begin{tabular}{l} 
meat, potatoes \\
and sweets
\end{tabular} & \(1.05(0.83,1.32)\) & \(45.71(0.056)\) & obese \\
\hline allergic rhinitis & \begin{tabular}{l} 
fruit, fish and \\
vegetables
\end{tabular} & \(0.97(0.84,1.12)\) & \(21.95(0.241)\) & obese \\
\hline allergic rhinitis & \begin{tabular}{l} 
meat, potatoes \\
and sweets
\end{tabular} & \(1.09(0.97,1.23)\) & \(0.00(0.480)\) & obese \\
\hline eczema & \begin{tabular}{l} 
fruit, fish and \\
vegetables
\end{tabular} & \(1.04(0.87,1.24)\) & \(46.01(0.054)\) & obese \\
\hline eczema & \begin{tabular}{l} 
meat, potatoes \\
and sweets
\end{tabular} & \(1.01(0.88,1.15)\) & \(15.51(0.300)\) & obese \\
\hline
\end{tabular}

\section*{XIII. Associations between specific food items that were declared significant from adjusted ESFA (FDR \(=5 \%\) ) and respiratory outcomes results of random intercept models; stratified by smoking status.}
\begin{tabular}{|c|c|c|c|c|}
\hline strata & Food item & Respiratory outcome & odds ratio & \(p\)-value \\
\hline non smoker & kindey and black beans & asthma & 0.96 & 0.562 \\
\hline ever smoker & kindey and black beans & asthma & 0.63 & 0.007 \\
\hline current smoker & kindey and black beans & asthma & 0.91 & 0.060 \\
\hline non smoker & lemon & asthma & 0.78 & 0.017 \\
\hline ever smoker & lemon & asthma & 0.87 & 0.177 \\
\hline current smoker & lemon & asthma & 0.85 & 0.017 \\
\hline non smoker & smoked fatty fish & asthma & 0.89 & 0.021 \\
\hline ever smoker & smoked fatty fish & asthma & 0.77 & 0.004 \\
\hline current smoker & smoked fatty fish & asthma & 0.84 & 0.018 \\
\hline non smoker & cherry & asthma & 0.85 & 0.000 \\
\hline ever smoker & cherry & asthma & 0.84 & 0.084 \\
\hline current smoker & cherry & asthma & 0.88 & 0.198 \\
\hline non smoker & any legumes & asthma & 1.00 & 0.911 \\
\hline ever smoker & any legumes & asthma & 0.74 & 0.021 \\
\hline current smoker & any legumes & asthma & 0.89 & 0.366 \\
\hline non smoker & condensed milk & asthma & 0.85 & 0.014 \\
\hline ever smoker & condensed milk & asthma & 0.92 & 0.016 \\
\hline current smoker & condensed milk & asthma & 0.88 & 0.019 \\
\hline non smoker & thin biscuits & asthma & 1.10 & 0.000 \\
\hline ever smoker & thin biscuits & asthma & 1.11 & 0.047 \\
\hline current smoker & thin biscuits & asthma & 1.03 & 0.670 \\
\hline non smoker & turnip & asthma & 1.14 & 0.069 \\
\hline ever smoker & turnip & asthma & 1.19 & 0.087 \\
\hline current smoker & turnip & asthma & 1.01 & 0.947 \\
\hline non smoker & couscous & asthma & 0.97 & 0.568 \\
\hline ever smoker & couscous & asthma & 1.31 & 0.000 \\
\hline current smoker & couscous & asthma & 0.96 & 0.849 \\
\hline non smoker & cabbage & chronic sinusitis & 0.91 & 0.026 \\
\hline ever smoker & cabbage & chronic sinusitis & 0.82 & 0.151 \\
\hline current smoker & cabbage & chronic sinusitis & 0.85 & 0.213 \\
\hline non smoker & viili or fermented yogurt & chronic sinusitis & 1.21 & 0.063 \\
\hline ever smoker & viili or fermented yogurt & chronic sinusitis & 1.20 & 0.010 \\
\hline current smoker & viili or fermented yogurt & chronic sinusitis & 1.12 & 0.208 \\
\hline non smoker & okra & chronic sinusitis & 1.34 & 0.000 \\
\hline ever smoker & okra & chronic sinusitis & 1.16 & 0.024 \\
\hline current smoker & okra & chronic sinusitis & 1.06 & 0.243 \\
\hline non smoker & pumpkin & chronic sinusitis & 1.20 & 0.001 \\
\hline
\end{tabular}
\begin{tabular}{|c|c|c|c|c|}
\hline ever smoker & pumpkin & chronic sinusitis & 1.21 & 0.036 \\
\hline current smoker & pumpkin & chronic sinusitis & 1.02 & 0.928 \\
\hline non smoker & peach & allergic rhinitis & 0.87 & 0.055 \\
\hline ever smoker & peach & allergic rhinitis & 0.81 & 0.001 \\
\hline current smoker & peach & allergic rhinitis & 0.60 & 0.000 \\
\hline non smoker & peanuts & allergic rhinitis & 0.77 & 0.338 \\
\hline ever smoker & peanuts & allergic rhinitis & 0.97 & 0.682 \\
\hline current smoker & peanuts & allergic rhinitis & 0.95 & 0.574 \\
\hline non smoker & tofu & allergic rhinitis & 0.80 & 0.013 \\
\hline ever smoker & tofu & allergic rhinitis & 0.99 & 0.805 \\
\hline current smoker & tofu & allergic rhinitis & 1.15 & 0.233 \\
\hline non smoker & mango & allergic rhinitis & 1.05 & 0.472 \\
\hline ever smoker & mango & allergic rhinitis & 1.07 & 0.450 \\
\hline current smoker & mango & allergic rhinitis & 1.35 & 0.049 \\
\hline non smoker & vegetable oil & allergic rhinitis & 1.21 & 0.002 \\
\hline ever smoker & vegetable oil & allergic rhinitis & 1.10 & 0.144 \\
\hline current smoker & vegetable oil & allergic rhinitis & 1.14 & 0.149 \\
\hline non smoker & smoked poultry & allergic rhinitis & 1.17 & 0.000 \\
\hline ever smoker & smoked poultry & allergic rhinitis & 0.91 & 0.078 \\
\hline current smoker & smoked poultry & allergic rhinitis & 1.17 & 0.000 \\
\hline non smoker & single cream & allergic rhinitis & 1.28 & 0.000 \\
\hline ever smoker & single cream & allergic rhinitis & 1.22 & 0.201 \\
\hline current smoker & single cream & allergic rhinitis & 1.31 & 0.139 \\
\hline non smoker & rhubarb & eczema & 0.86 & 0.005 \\
\hline ever smoker & rhubarb & eczema & 0.85 & 0.008 \\
\hline current smoker & rhubarb & eczema & 0.69 & 0.045 \\
\hline non smoker & crisp fried Cakes & eczema & 0.86 & 0.105 \\
\hline ever smoker & crisp fried Cakes & eczema & 0.82 & 0.024 \\
\hline current smoker & crisp fried Cakes & eczema & 0.76 & 0.060 \\
\hline non smoker & bitter mellon & eczema & 0.78 & 0.001 \\
\hline ever smoker & bitter mellon & eczema & 0.95 & 0.564 \\
\hline current smoker & bitter mellon & eczema & 0.96 & 0.563 \\
\hline non smoker & greek style yogurt & eczema & 0.79 & 0.003 \\
\hline ever smoker & greek style yogurt & eczema & 0.85 & 0.000 \\
\hline current smoker & greek style yogurt & eczema & 1.26 & 0.212 \\
\hline non smoker & other game & eczema & 0.96 & 0.025 \\
\hline ever smoker & other game & eczema & 1.12 & 0.184 \\
\hline current smoker & other game & eczema & 0.76 & 0.005 \\
\hline non smoker & lentils & eczema & 1.06 & 0.049 \\
\hline ever smoker & lentils & eczema & 0.92 & 0.000 \\
\hline current smoker & Ientils & eczema & 0.76 & 0.000 \\
\hline non smoker & sour cream & eczema & 1.22 & 0.019 \\
\hline ever smoker & sour cream & eczema & 1.10 & 0.080 \\
\hline current smoker & sour cream & eczema & 0.83 & 0.292 \\
\hline
\end{tabular}

Appendices
\begin{tabular}{|l|l|l|l|l|}
\hline non smoker & crisp fried Cakes & atopy & 0.63 & 0.006 \\
\hline ever smoker & crisp fried Cakes & atopy & 1.01 & 0.858 \\
\hline current smoker & crisp fried Cakes & atopy & 0.92 & 0.278 \\
\hline non smoker & moussaka & atopy & 0.92 & 0.212 \\
\hline ever smoker & moussaka & atopy & 0.97 & 0.789 \\
\hline current smoker & moussaka & atopy & 0.53 & 0.000 \\
\hline non smoker & beer & atopy & 1.16 & 0.148 \\
\hline ever smoker & beer & atopy & 1.08 & 0.388 \\
\hline current smoker & beer & atopy & 1.14 & 0.119 \\
\hline non smoker & thin biscuits & atopy & \(\mathbf{1 . 2 3}\) & \(\mathbf{0 . 0 0 0}\) \\
\hline ever smoker & thin biscuits & atopy & \(\mathbf{0 . 8 9}\) & \(\mathbf{0 . 0 2 1}\) \\
\hline current smoker & thin biscuits & atopy & \(\mathbf{1 . 1 9}\) & \(\mathbf{0 . 0 0 1}\) \\
\hline non smoker & roast beef steak & atopy & 1.24 & 0.011 \\
\hline ever smoker & roast beef steak & atopy & 1.11 & 0.174 \\
\hline current smoker & roast beef steak & 1.07 & 0.350 \\
\hline
\end{tabular}

\section*{XIV. Associations between specific food items that were declared significant from adjusted ESFA (FDR \(\mathbf{= 5 \%}\) ) and respiratory outcomes; results of random intercept models; stratified by gender.}
\begin{tabular}{|c|c|c|c|c|}
\hline strata & food item & respiratory outcome & odds ratio & pvalue \\
\hline male & kidney and black beans & asthma & 0.85 & 0.003 \\
\hline female & kidney and black beans & asthma & 0.79 & 0.004 \\
\hline male & lemon & asthma & 0.87 & 0.068 \\
\hline female & lemon & asthma & 0.84 & 0.000 \\
\hline male & smoked fatty fish & asthma & 0.82 & 0.000 \\
\hline female & smoked fatty fish & asthma & 0.87 & 0.000 \\
\hline male & cherry & asthma & 0.71 & 0.160 \\
\hline female & cherry & asthma & 0.89 & 0.000 \\
\hline male & any legumes & asthma & 0.80 & 0.000 \\
\hline female & any legumes & asthma & 0.95 & 0.096 \\
\hline male & condensed milk & asthma & 0.80 & 0.024 \\
\hline female & condensed milk & asthma & 0.93 & 0.026 \\
\hline male & thin biscuits & asthma & 1.20 & 0.000 \\
\hline female & thin biscuits & asthma & 1.01 & 0.747 \\
\hline male & turnsip & asthma & 1.09 & 0.380 \\
\hline female & turnsip & asthma & 1.15 & 0.056 \\
\hline male & couscous & asthma & 1.19 & 0.000 \\
\hline female & couscous & asthma & 1.12 & 0.128 \\
\hline male & cabbage & chronic sinusitis & 0.69 & 0.022 \\
\hline female & cabbage & chronic sinusitis & 0.90 & 0.051 \\
\hline male & viili and fermented yogurt & chronic sinusitis & 0.94 & 0.476 \\
\hline female & viili and fermented yogurt & chronic sinusitis & 1.30 & 0.000 \\
\hline male & okra & chronic sinusitis & 1.01 & 0.756 \\
\hline female & okra & chronic sinusitis & 1.19 & 0.154 \\
\hline male & pumpkin & chronic sinusitis & 1.29 & 0.063 \\
\hline female & pumpkin & chronic sinusitis & 1.16 & 0.000 \\
\hline male & peach & allergic rhinitis & 0.87 & 0.183 \\
\hline female & peach & allergic rhinitis & 0.81 & 0.001 \\
\hline male & peanuts & allergic rhinitis & 0.90 & 0.079 \\
\hline female & peanuts & allergic rhinitis & 0.91 & 0.390 \\
\hline male & tofu & allergic rhinitis & 0.91 & 0.023 \\
\hline female & tofu & allergic rhinitis & 0.81 & 0.016 \\
\hline male & mango & allergic rhinitis & 1.15 & 0.391 \\
\hline female & mango & allergic rhinitis & 1.10 & 0.002 \\
\hline male & vegetable oil & allergic rhinitis & 1.22 & 0.000 \\
\hline female & vegetable oil & allergic rhinitis & 1.09 & 0.134 \\
\hline
\end{tabular}
\begin{tabular}{|l|l|l|r|r|}
\hline male & smoked poultry & allergic rhinitis & 1.08 & 0.038 \\
\hline female & smoked poultry & allergic rhinitis & 1.12 & 0.000 \\
\hline male & single cream & allergic rhinitis & 1.15 & 0.201 \\
\hline female & single cream & allergic rhinitis & 1.46 & 0.013 \\
\hline male & rhubarb & eczema & 0.86 & 0.127 \\
\hline female & rhubarb & eczema & 0.84 & 0.000 \\
\hline male & crisp fried Cakes & eczema & 0.81 & 0.036 \\
\hline female & crisp fried Cakes & eczema & 0.86 & 0.040 \\
\hline male & bitter mellon & eczema & 0.86 & 0.176 \\
\hline female & bitter mellon & eczema & 0.90 & 0.004 \\
\hline male & greek style yogurt & eczema & 1.03 & 0.802 \\
\hline female & greek style yogurt & eczema & 0.82 & 0.001 \\
\hline male & other game & eczema & 1.07 & 0.115 \\
\hline female & other game & eczema & 0.88 & 0.000 \\
\hline male & lentils & eczema & 0.89 & 0.000 \\
\hline female & lentils & eczema & 1.01 & 0.495 \\
\hline male & sour cream & eczema & 1.07 & 0.171 \\
\hline female & sour cream & atopy & 1.20 & 0.004 \\
\hline male & crisp fried Cakes & atopy & 0.97 & 0.619 \\
\hline female & crisp fried Cakes & atopy & 0.72 & 0.031 \\
\hline male & moussaka & atopy & \(\mathbf{0 . 8 9}\) & \(\mathbf{0 . 0 0 4}\) \\
\hline female & moussaka & atopy & \(\mathbf{0 . 8 8}\) & \(\mathbf{0 . 0 0 6}\) \\
\hline male & beer & atopy & \(\mathbf{1 . 0 9}\) & \(\mathbf{0 . 0 0 2}\) \\
\hline female & beer & atopy & \(\mathbf{1 . 3 2}\) & \(\mathbf{0 . 0 3 8}\) \\
\hline male & thin biscuits & atopy & \(\mathbf{0 . 0 5 1}\) \\
\hline female & thin biscuits & atopy & \(\mathbf{0 . 0 0 1}\) \\
\hline male & roast beef steak & \(\mathbf{0 . 0 2 5}\) \\
\hline female & roast beef steak & \(\mathbf{0 . 0 3 4}\) \\
\hline & & & 0.00 \\
\hline
\end{tabular}

\section*{XV. Associations between specific food items that were declared significant from adjusted ESFA (FDR \(=5 \%\) ) and respiratory outcomes; results of random intercept models; stratified by body mass index.}
\begin{tabular}{|c|c|c|c|c|}
\hline strata & respiratory outcome & food item & odds ratio & pvalue \\
\hline non-obese & asthma & kidney and black beans & 0.79 & 0.000 \\
\hline obese & asthma & kidney and black beans & 0.83 & 0.000 \\
\hline non-obese & asthma & lemon & 0.84 & 0.046 \\
\hline obese & asthma & lemon & 0.85 & 0.000 \\
\hline non-obese & asthma & smoked fatty fish & 0.98 & 0.781 \\
\hline obese & asthma & smoked fatty fish & 0.78 & 0.000 \\
\hline non-obese & asthma & cherry & 0.86 & 0.314 \\
\hline obese & asthma & cherry & 0.83 & 0.004 \\
\hline non-obese & asthma & any legumes & 0.88 & 0.251 \\
\hline obese & asthma & any legumes & 0.85 & 0.012 \\
\hline non-obese & asthma & condensed milk & 0.97 & 0.725 \\
\hline obese & asthma & condensed milk & 0.85 & 0.071 \\
\hline non-obese & asthma & thin biscuits & 1.05 & 0.003 \\
\hline obese & asthma & thin biscuits & 1.16 & 0.000 \\
\hline non-obese & asthma & turnsip & 1.07 & 0.354 \\
\hline obese & asthma & turnsip & 1.20 & 0.083 \\
\hline non-obese & asthma & couscous & 1.08 & 0.446 \\
\hline obese & asthma & couscous & 1.20 & 0.000 \\
\hline non-obese & chronic sinusitis & cabbage & 0.86 & 0.517 \\
\hline obese & chronic sinusitis & cabbage & 0.90 & 0.308 \\
\hline non-obese & chronic sinusitis & viili and fermented yogurt & 1.37 & 0.000 \\
\hline obese & chronic sinusitis & viili and fermented yogurt & 1.01 & 0.871 \\
\hline non-obese & chronic sinusitis & okra & 1.21 & 0.001 \\
\hline obese & chronic sinusitis & okra & 1.06 & 0.096 \\
\hline non-obese & chronic sinusitis & pumpkin & 1.16 & 0.000 \\
\hline obese & chronic sinusitis & pumpkin & 1.19 & 0.006 \\
\hline non-obese & allergic rhinitis & peach & 0.69 & 0.038 \\
\hline obese & allergic rhinitis & peach & 0.93 & 0.455 \\
\hline non-obese & allergic rhinitis & peanuts & 0.94 & 0.259 \\
\hline obese & allergic rhinitis & peanuts & 0.89 & 0.001 \\
\hline non-obese & allergic rhinitis & tofu & 0.96 & 0.441 \\
\hline obese & allergic rhinitis & tofu & 0.79 & 0.302 \\
\hline non-obese & allergic rhinitis & mango & 1.27 & 0.198 \\
\hline obese & allergic rhinitis & mango & 0.99 & 0.950 \\
\hline non-obese & allergic rhinitis & vegetable oil & 1.19 & 0.001 \\
\hline obese & allergic rhinitis & vegetable oil & 1.12 & 0.290 \\
\hline non-obese & allergic rhinitis & smoked poultry & 0.97 & 0.441 \\
\hline obese & allergic rhinitis & smoked poultry & 1.24 & 0.019 \\
\hline non-obese & allergic rhinitis & single cream & 1.22 & 0.150 \\
\hline
\end{tabular}
\begin{tabular}{|l|l|l|l|l|}
\hline obese & allergic rhinitis & single cream & 1.32 & 0.000 \\
\hline non-obese & eczema & rhubarb & 0.69 & 0.000 \\
\hline obese & eczema & rhubarb & 0.94 & 0.270 \\
\hline non-obese & eczema & crisp fried Cakes & 0.84 & 0.003 \\
\hline obese & eczema & crisp fried Cakes & 0.86 & 0.108 \\
\hline non-obese & eczema & bitter mellon & \(\mathbf{0 . 8 8}\) & \(\mathbf{0 . 0 4 8}\) \\
\hline obese & eczema & bitter mellon & 0.89 & \(\mathbf{0 . 0 0 0}\) \\
\hline non-obese & eczema & greek style yogurt & 0.99 & 0.041 \\
\hline obese & eczema & greek style yogurt & 0.89 & 0.077 \\
\hline non-obese & eczema & other game & 1.05 & 0.001 \\
\hline obese & eczema & other game & 0.93 & \(\mathbf{0 . 9 5}\) \\
\hline non-obese & eczema & lentils & \(\mathbf{0 . 0 0 0}\) \\
\hline obese & eczema & lentils & 1.09 & \(\mathbf{0 . 0 4 7}\) \\
\hline non-obese & eczema & sour cream & 0.161 \\
\hline obese & eczema & crisp fried Cakes & 0.76 & 0.000 \\
\hline non-obese & atopy & crisp fried Cakes & 0.93 & 0.027 \\
\hline obese & atopy & moussaka & 0.209 \\
\hline non-obese & atopy & moussaka & 0.91 & 0.026 \\
\hline obese & atopy & beer & 1.09 & 0.074 \\
\hline non-obese & atopy & beer & 1.12 & 0.116 \\
\hline obese & atopy & thin biscuits & 1.31 & 0.084 \\
\hline non-obese & atopy & thin biscuits & 0.000 \\
\hline obese & atopy & roast beef steak & 0.290 \\
\hline non-obese & atopy & roast beef steak & 1.26 & 0.010 \\
\hline obese & atopy & & 0.165 \\
\hline & & 0 & \\
\hline
\end{tabular}```


[^0]:    ${ }^{\text {a }}$ Data-set 1 is from the FLAG survey (Shaheen et al., 2001); Data-set 2 is from the UK ECRHS II survey (Hooper et al., 2010). In Model 1 all selected food intakes are negatively associated with disease; in Model 2 half the selected food intakes are negatively associated and half are positively associated; and in Model 3 all selected food intakes are positively associated with disease. ${ }^{\text {b }}$ Power of PCA exceeds that of ESFA.

[^1]:    ${ }^{\text {a }}$ Data-set 1 is from the FLAG survey (Shaheen et al., 2001); Data-set 2 is from the UK ECRHS II survey (Hooper et al., 2010). In Model 1 all selected food intakes are negatively associated with disease; in Model 2 half the selected food intakes are negatively associated and half are positively associated; and in Model 3 all selected food intakes are positively associated with disease. ${ }^{\mathrm{b}}$ Power of PCA exceeds that of ESFA, but FDR is also higher. ${ }^{\mathrm{c}}$ Power is defined as the number of true significant results identified by the method divided by the number of foods that are causally linked with disease. ${ }^{\mathrm{d}} \mathbf{F D R}$ is defined as the number of false significant results identified by the method divided by the total number of significant results identified by the method. FDR of Simes procedure at $20 \%$.

[^2]:    ${ }^{\text {a }}$ Data-set 1 is from the FLAG survey (Shaheen et al., 2001); Data-set 2 is from the UK ECRHS II survey (Hooper et al., 2010). In Model 1 all selected food intakes are negatively associated with disease; in Model 2 half the selected food intakes are negatively associated and half are positively associated; and in Model 3 all selected food intakes are positively associated with disease. ${ }^{\text {b }}$ Power of PCA exceeds that of ESFA, but FDR is also higher. ${ }^{\text {c Power is defined as the number of true }}$ significant results identified by the method divided by the number of foods that are causally linked with disease. ${ }^{\mathbf{d}} \mathbf{F D R}$ is defined as the number of false significant results identified by the method divided by the total number of significant results identified by the method. FDR of Simes procedure at $20 \%$.

[^3]:    ${ }^{\text {a }}$ Data-set 1 is from the FLAG survey (Shaheen et al., 2001); Data-set 2 is from the UK ECRHS II survey (Hooper et al., 2010). In Model 4, foods being included in a "Western" pattern ( 30 in Data-set 1 and 10 in Data-set 2 - see paragraph 4.9.2 and 4.9.3) are used in each replication, with pattern intake being positively associated with disease. In Model 5 , foods being included in a "Western" pattern ( 30 in Data-set 1 and 10 in Data-set 2 - see paragraph 4.9.2 and 4.9.3) are used in each replication, with pattern intake being negatively associated. ${ }^{\text {c }}$ FDR of PCA is lower than that of ESFA, but power is also lower. ${ }^{\text {c }}$ Power is defined as the number of true significant results identified by the method divided by the number of foods that are causally linked with disease. ${ }^{d}$ FDR is defined as the number of false significant results identified by the method divided by the total number of significant results identified by the method. FDR of Simes procedure at $20 \%$.

[^4]:    ${ }^{\text {a }}$ Data-set 1 is from the FLAG survey (Shaheen et al., 2001); Data-set 2 is from the UK ECRHS II survey (Hooper et al., 2010). In Model 1 all selected food intakes are negatively associated with disease; in Model 2 half the selected food intakes are negatively associated and half are positively associated; and in Model 3 all selected food intakes are positively associated with disease. ${ }^{\text {c }}$ Power is defined as the number of true significant results identified by the method divided by the number of foods that are causally linked with disease. ${ }^{\mathrm{d}}$ FDR is defined as the number of false significant results identified by the method divided by the total number of significant results identified by the method. FDR of Simes procedure at $20 \%$.

[^5]:    ${ }^{\text {a }}$ Data-set 1 is from the FLAG survey (Shaheen et al., 2001); Data-set 2 is from the UK ECRHS II survey (Hooper et al., 2010). In Model 1 all selected food intakes are negatively associated with disease; in Model 2 half the selected food intakes are negatively associated and half are positively associated; and in Model 3 all selected food intakes are positively associated with disease. ${ }^{\text {c }}$ Power is defined as the number of true significant results identified by the method divided by the number of foods that are causally linked with disease. ${ }^{\mathbf{d}} \mathbf{F D R}$ is defined as the number of false significant results identified by the method divided by the total number of significant results identified by the method. FDR of Simes procedure at $20 \%$.

[^6]:    ${ }^{\text {a }}$ Data-set 1 is from the FLAG survey (Shaheen et al., 2001); Data-set 2 is from the UK ECRHS II survey (Hooper et al., 2010). In Model 4, foods being included in a "Western" pattern (30 in Data-set 1 and 10 in Data-set 2 - see paragraph 4.9.2 and 4.9.3) are used in each replication, with pattern intake being positively associated with disease. In Model 5, foods being included in a "Western" pattern (30 in Data-set 1 and 10 in Data-set 2 - see paragraph 4.9.2 and 4.9.3) are used in each replication, with pattern intake being negatively associated. ${ }^{\text {c }}$ Power is defined as the number of true significant results identified by the method divided by the number of foods that are causally linked with disease. ${ }^{\mathbf{d}} \mathbf{F D R}$ is defined as the number of false significant results identified by the method divided by the total number of significant results identified by the method. FDR of Simes procedure at $20 \%$.

[^7]:    *Adjusted for age, sex, body mass index and smoking status. **Adjusted for age, sex, body mass index, smoking status and all the foods that were significant at the univariate analysis ( $1^{\text {st }}$ step). ***Bold type indicates foods which were statistically significant associated with eczema across 3 or more of the different procedures being used.

