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1	A tilt after-effect for images of buildings:
2	Evidence of selectivity for the orientation of everyday scenes
3	
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Abstract

21

The tilt after-effect (TAE) is thought to be a manifestation of gain control in mechanisms 22 selective for spatial orientation in visual stimuli. It has been demonstrated with luminance-23 24 defined stripes, contrast-defined stripes, orientation-defined stripes, and even with natural images. Of course, all images can be decomposed into a sum of stripes, so it should not be 25 surprising to find a TAE when adapting and test images contain stripes that differ by 15° or 26 so. We show this latter condition is not necessary for the TAE with natural images: 27 adaptation to slightly tilted and vertically filtered houses produced a "repulsive" bias in the 28 perceived orientation of horizontally filtered houses. These results suggest gain control in 29 mechanisms selective for spatial orientation in natural images. 30

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33 Keywords: natural images, global orientation, tilt after-effect, spatially non-specific

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Introduction

Gibson and Radner [1] demonstrated that adapting to a line tilted between 2.5° and 45° from 37 vertical makes a vertical "test" stimulus, presented in the same retinal location, appear tilted 38 in a direction opposite to that of the adaptor. This repulsive effect on perceived orientation is 39 known as the tilt after-effect (TAE). Most contemporary theories commonly attribute the 40 TAE to suppression of responses in neurons tuned to the adaptor's orientation [2], either via 41 fatigue of the adapted neurons [3] or lateral inhibition between neurons with similar 42 orientation preferences [4, 5], although other accounts have been proposed [6]. The TAE is a 43 44 natural consequence of orientation-selective suppression, which effectively skews neural responses away from the adapting orientation. 45

46

47 Any repulsive after-effect can be considered as evidence for the existence of neural populations selectively encoding a specific stimulus feature. Consequently, after-effects have 48 earned a reputation for being "the psychophysicist's micro-electrode" [7]. Using after-effects, 49 50 psychophysicists have inferred the existence of neural selectivity for such complex attributes as shape, glossiness, and facial expression [8]. There is even an after-effect of adaptation to 51 heavily masculine or feminine features [9]. However, it must be acknowledged that some of 52 these after-effects might be the result of adaptation in "low-level" visual mechanisms, tuned 53 to stimulus values that have nothing to do with faces per se. For example, if adapting to a 54 55 thick, masculine eyebrow suppresses a few neurons that prefer (low spatial frequency) shapes like that, then a subsequently viewed, androgynous eyebrow (with a slightly higher spatial 56 frequency) will appear much thinner, making the face it is on appear more feminine. Thus, 57 58 inferring neural mechanisms from perceptual after-effects is not always as straightforward as one might hope. 59

Inferring neural selectivity from psychophysics is complicated, not only because after-effects 61 can reflect adaptation by low-level mechanisms, but also because many conventional 62 measurements of appearance are susceptible to contamination from non-perceptual sources of 63 bias (e.g., expectation effects and response biases; [10]). In this study, we minimize the 64 influence of low-level adaptation by restricting adaptor and tests to different regions of the 65 visual field and / or different regions of frequency space. We minimize the influence of non-66 perceptual sources of bias by adopting the recently developed, two-alternative, forced-choice 67 (2AFC) comparison-of-comparisons paradigm, with roving pedestals [11, 12]. 68

69

The after-effect we have studied is the recently reported TAE for natural scenes [13]. Global 70 scene orientation is important for a number of reasons. Firstly, perceived orientation of a 71 72 scene provides information about the direction of gravity, which in turn informs selforientation relative to gravity. This is particularly relevant when information provided by 73 other sensory sources is discordant [14]. Secondly, judgements of subjective visual vertical 74 75 are affected by the orientation of background scenes, which serve as a global frame of reference for perceptual judgements [15, 16]. Finally, it has been reported that scene 76 orientation affects how people deploy overt attention within a scene, where scene-centric 77 directional asymmetries of eye movements always remain aligned with the orientation of the 78 scene [17]. 79

80

In Experiment 1 we confirm that the TAE for natural scenes can be obtained with different (and differently sized) adapting and test images, which are presented in a partially overlapping spatial configuration and share minimal spatial frequency components. In Experiment 2, the specific question we address is whether the TAE for natural scenes arises because of interactions between mechanisms selective for natural scenes, or whether it is

simply a by-product of suppression between more lower-level mechanisms, selective for
spatial orientation in general. To disentangle these possibilities, we use orientation-filtered
and phase-scrambled stimuli. Vertically filtered images are designed to have a negligible
effect on the responsivity of low-level mechanisms tuned to near-horizontal orientations.
Phase-scrambled stimuli are designed to have a similarly negligible effect on the responsivity
of mechanisms selective for natural scenes.

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Methods

94 **Participants**

A total of 23 observers (18 – 46 years of age), each having a unique two-character set of initials (see figures 2 and 3), from Queen Mary University of London with normal or corrected-to-normal visual acuity took part in the experiments. Procedures were approved by the Queen Mary University of London research ethics committee and written informed consent was obtained from all participants. The number of participants for each experimental condition was determined based on previous studies investigating higher-level visual aftereffects, which involved from 5 to 10 observers per condition [18-20].

102

103 Experimental set-up and apparatus

Observers were seated in a dark room, and were instructed to keep their head upright and maintain the same distance from the screen throughout the experiment. Stimuli were presented on a 20" Iiyama CRT monitor with a 1600×1200 screen resolution and a refresh rate of 60 Hz. The viewing distance was 57 cm, such that each pixel subtended 1.5 arcminutes. A black mask with a circular aperture (diameter = 24.5°) was overlaid on the monitor to eliminate the use of monitor edges as cues to vertical or horizontal. Stimulus presentation and data collection used Matlab (Mathworks) and Psychtoolbox [21]. 111

112 Stimuli

Images of 5 different houses (figure 1B), in their frontal views, appearing to be at eye level 113 from a standing position, were obtained from an archive of the Caltech Computational Vision 114 Group (available online at http://www.vision.caltech.edu/archive.html). We used images of 115 houses because: 1) scene orientation of man-made scenes is judged with better discrimination 116 precision than non-man-made scenes [16] and 2) houses have a clear frontal facade and cover 117 limited depth, resulting in minimal linear perspectives. The images were initially cropped to a 118 119 square aspect ratio and then resized to 300×300 pixels using bicubic interpolation. Cropped images were converted to grayscale by independently weighting and summing the red, green 120 and blue channels of the image according to the CIE procedure (0.299 \times R + 0.587 \times G + 121 $0.114 \times B$). These images were presented as adaptors within a hard-edged circular aperture 122 (diameter = 7.5° ; figure 1A). The test images were resized to 75% of the adaptor's size and 123 presented within a hard edged window of diameter 5.7°. 124

125

Images of houses were tilted and, in some cases, filtered. Filtering was a 7-step procedure. In 126 step 1 the mean graylevel of a tilted image was subtracted, creating a difference image with 127 no DC component. In step 2 this difference image was multiplied with a 2-dimensional, 128 separable cosine window of the same size. In step 3 the windowed image was Fourier 129 130 transformed (applying the cosine window before Fourier transformation helps to reduce wrap-around artefacts). In step 4 the transformed image was multiplied by one of the filters 131 described below. In step 5 the product was inverse-Fourier transformed. In step 6 the image 132 133 was scaled such that adaptors would have a root mean square (RMS) contrast of 0.10 and tests would have an RMS contrast of 0.18. Finally, in step 7, a graylevel of 0.50 was added to 134 each image. This matched the graylevel of the screen background. 135

136

137 **Procedure**

Trials were blocked by condition (there were three conditions in Experiment 1 and two conditions in Experiment 2) and adaptor orientation: either -15° or $+15^{\circ}$. By convention, we consider tilts clockwise (CW) from vertical to be negative and tilts counter-clockwise (CCW) from vertical to be positive. Each condition in Experiment 1 and 2 was also associated with a "baseline block," in which no adaptor was shown.

143

144 The general procedure is outlined in figure 1A. Observers were instructed to fixate a centrally presented white circle (diameter = 0.2°) for the duration of each block. All blocks (except 145 baseline blocks) began with an initial adaptation phase of 20 s. Following this, each test trial 146 started with a "top-up" adaptation phase of 5 s. During adaptation phases, the adaptor was 147 jittered every 0.5 s by recentering it on a random pixel within a predefined jitter area of 0.25° 148 \times 0.25° surrounding fixation. Top-up adaptors were followed, after 0.25 s, by two test houses, 149 presented immediately to the left and right of fixation, for 0.05 s. One of the test houses was 150 the "pedestal," with one of two fixed tilts: -3° or $+3^{\circ}$. The other test was the "comparison," 151 with an offset added to the fixed tilt, randomly selected from the set $\{-15^\circ, -12^\circ, -9^\circ, -6^\circ, -12^\circ, -9^\circ, -9^\circ, -6^\circ, -12^\circ, -9^\circ, -6^\circ, -12^\circ, -9^\circ, -9$ 152 3° , 0° , $+3^{\circ}$, $+6^{\circ}$, $+9^{\circ}$, $+12^{\circ}$, $+15^{\circ}$ }. Each combination of pedestal and comparison tilt was 153 tested 10 times, resulting in 220 trials per block. The spatial positions (left and right of 154 155 fixation) of the pedestal and comparison were randomized on every trial. Observers chose which of the two test houses appeared more upright, using keys "1" (for left) and "2" (for 156 right). Observers were told that an upright house is how they would imagine it to appear, if 157 they stood in front of it with their head held straight. 158

As is evident from figure 1A, there was a small amount of spatial overlap between the adaptor and tests. However, the overlapping parts of the images were not the same (e.g., the right half of the adaptor overlapped with the left half of one test) and were of different sizes to reduce retinotopic adaptation [22].

164

165 Methods specific to Experiment 1

In the *same house* condition image H1 was used for both adaptor and test stimuli. In the *different house* condition image H2 was the adaptor and image H3 was used for the tests (figure 1B). In the *different SF house* condition the adaptor and test stimuli were images of the same house, but filtered to separate them for their spatial frequency (SF) content (figure 2B). In this condition, three different house images were used (H2, H4 & H5; figure 1B). Two observers were tested with H2, two with H4 and two with H5.

172

Log-normal filters were used for the *different SF house* condition. The filter used for adaptors
had a peak SF of 10 cycles / degree. The filter used for the tests had a peak SF of 1.25 cycles
/ degree. Both filters had a half-bandwidth at half-height of 1.5 octaves.

176

177 Methods specific to Experiment 2

All 10 observers participated in both the *orthogonal house* condition and the *phasescrambled house* condition. In both conditions adaptors were first tilted (either CW or CCW) and then filtered to retain Fourier energy close to vertical orientations (figure 3). Tests were upright images of the same house, initially filtered horizontally and then tilted by different amounts in each trial, as in Experiment 1. Five observers were tested using H1; the other five were tested using H2. For each observer, the adapting and test stimuli were differently filtered versions of the same house image. In the orientation domain, each filter was a

Gaussian function of angle, centred on 0° (for the vertically filtered adaptors) or 90° (for the 185 horizontally filtered tests); with a half-bandwidth at half-height of 23.5° and was clipped at \pm 186 40° from the peak, resulting in zero gain at orientations beyond the clip. In the phase-187 scrambled condition, tilted adaptors were phase-scrambled prior to orientation filtering, by 188 adding a uniform distribution of random phase offsets (between $-\pi$ and $+\pi$) to the Fourier 189 phases of the image. The power spectra and RMS contrast of adaptors in the *phase-scrambled* 190 191 house condition matched the power spectra and RMS contrast of adaptors in the orthogonal house condition. Identical (unscrambled), horizontally filtered, tilted tests were used in both 192 193 conditions.

194

195 Psychophysical model

196 Data were analysed within the context of signal-detection theory, as described by Morgan, Grant, Melmoth, & Solomon [23]. Within this model, the appearances of pedestal (S) and 197 comparison (C) are normally distributed, i.e., $S \sim N(p + \mu, \sigma^2/2)$ and $C \sim N(p + \mu + \mu)$ 198 $t, \sigma^2/2$), where σ^2 is the variance of the performance-limiting noise, p is the pedestal tilt, t is 199 the offset added to the comparison, and μ is the perceptual bias specific to each test block. If 200 there were no perceptual bias, then the distributions for pedestal and comparison would have 201 means of p and p + t respectively. The observer chooses the pedestal as closer to upright 202 when it appears less tilted than the comparison. Accordingly, the probability of this choice 203 $P("S") = P(|S| < |C|) = P(S^2/C^2 < 1)$, has a doubly non-central F distribution. This 204 distribution's denominator's noncentrality parameter is $2(p + \mu + t)^2 / \sigma^2$, its numerator's 205 noncentrality parameter is $2(p + \mu)^2 / \sigma^2$, and both denominator and numerator have 1 206 degree of freedom. 207

208

Results

210

From each block of trials (baseline, CCW and CW), we obtained maximum-likelihood estimates of bias μ and the variance of performance-limiting noise σ^2 . Negative biases with CCW adaptors and positive biases with CW adaptors are indicative of the repulsive TAE. Non-parametric bootstrapping (with bias-correction [24]) was used to quantify the reliability of our parameter estimates. The error bars shown in figures 2 and 3 contain the resultant 95% confidence intervals.

217

We also fit each observer's data from CCW-adaptor and CW-adaptor blocks simultaneously, 218 forcing the bias parameter μ to be the same in both cases, but allowing σ to vary. The ratio L, 219 between the likelihood of this nested model fit and the joint likelihood of the aforementioned 220 separate fits to the same data is necessarily no greater than 1. To evaluate the "null" 221 hypothesis of no significant TAE in individual observers, we compare the criteria $\alpha = 0.05$ 222 and $\alpha = 0.001$ to the value $1 - F(-2 \ln L)$, where F is the cumulative chi-square distribution, 223 with 1 degree of freedom. This is known as the generalized likelihood-ratio test (see [25] 224 p.440-441). 225

226

To evaluate null hypotheses at the group level, we performed one-sample *t*-tests using 227 estimates of repulsion, which can be quantified either in degrees of tilt or in terms of the 228 "just-noticeable difference" (JND). A single value for repulsion, in degrees of tilt, can be 229 obtained by subtracting one maximum-likelihood estimate of μ (the one obtained with CCW) 230 adaptors) from the complimentary estimate (obtained with CW adaptors), and dividing the 231 difference by 2. The "conspicuousness" of repulsion can be quantified by further dividing this 232 quotient by the JND. For the latter, we use the root-mean-square of the maximum-likelihood 233 estimates of σ . Results of the group-level *t*-tests appear in tables 1 and 2. 234

235

236 Experiment 1

Estimates of bias (μ) from Experiment 1 are plotted in figure 2A. For the majority of observers, adaptation to a house tilted 15° (CCW of upright) produced a negative bias (relative to the baseline's bias) in subsequently viewed test houses, and adaptation to a house tilted -15° produced a positive bias. Generalized likelihood ratio tests suggest after-effects significant at the $\alpha = 0.05$ level for repulsion in the data from 5 of the 7 observers in the *same house* condition, 5 of the 6 observers in the *different house* condition, and all 6 of the 6 observers in the *different SF house* condition. Group-level statistics appear in tables 1 and 2.

244

245 Experiment 2

Estimates of bias from Experiment 2 are plotted in figure 3. Generalized likelihood ratio tests 246 suggest after-effects significant at the $\alpha = 0.05$ level for repulsion in the data from 8 of the 10 247 observers in the orthogonal house condition and none of the (same) 10 observers in the 248 phase-scrambled house condition. Group-level statistics appear in tables 1 and 2. At the 249 250 group level, both conditions produced mean repulsion and conspicuousness significantly larger than zero. However, a comparison using a paired-samples *t*-test between the means of 251 the two conditions revealed that the *orthogonal house* condition produced a significantly 252 253 larger repulsion compared to the *phase-scrambled house* condition (tables 1 & 2).

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Discussion

Our results (Experiment 1) demonstrate that the TAE for natural scenes (houses) can be obtained with partially overlapping, yet different (and differently sized) adapting and test images, widely separated in spatial frequency content. Similar results have been obtained with sinusoidal gratings [18, 26] and circular / radial patterns [19]. When after-effects survive

manipulations of image, size and spatial frequency, their origin cannot be attributed to low-260 level visual mechanisms [22]. Our results extend Dekel & Sagi's [13] findings of TAEs with 261 natural images as adaptors and sinusoidal gratings as tests, by showing that adaptation to 262 global orientation can occur between adaptors and tests that are natural images. However, it is 263 unclear from Experiment 1 whether the TAE for natural scenes arises because of interactions 264 between high-level mechanisms selective for natural scenes, or whether it is simply a by-265 product of suppression between mid-level mechanisms, selective for spatial orientation in 266 general. 267

268

To distinguish between these alternatives, in Experiment 2 we applied perpendicular filters to 269 our stimuli, widely separating the orientation contents of adaptor and tests. Our finding of a 270 271 repulsive TAE in this condition qualitatively differs from the assimilative "indirect effect" found when retinally overlapping lines or gratings are separated between 60° and 87.5° [1]. 272 We attribute this repulsion to our images' recognisability as slightly tilted scenes, rather than 273 their Fourier image components. In support of this viewpoint, we found no after-effect at the 274 individual observer level when the Fourier phases of our adaptors were scrambled. However, 275 the group level analyses did reveal a relatively small but significant TAE (tables 1 & 2), with 276 phase-scrambled adaptors. This must be attributed to Fourier image components. A possible 277 reason for this is that since man-made images are usually dominated by cardinal orientations, 278 279 a sense of global tilt is still apparent in the images even after randomizing Fourier phase information (see figure 3B, where randomized images might appear tilted CW). 280

281

Our most interesting finding is that vertically filtered houses induce repulsive TAEs. These TAEs were not only evident in most observers, but they were also much larger than the TAEs from phase-scrambled adaptors at the group level. Although our orientation-filtered houses

are not as easily recognizable as their unfiltered counterparts, they possess clear higher-order 285 structure, which is lacking in the phase-scrambled versions used for adaptation. Textures with 286 similar higher-order (meaningless) structure are also more effective than phase-scrambled 287 scenes as backward masks of 'scene gist' [27]. This suggests that textures with higher-order 288 structure are fundamentally different from phase-randomized stimuli with similar orientation 289 statistics. Nonetheless, the after-effect of adapting to tilted buildings is different from the 290 after-effect elicited by the perception of a global form contained in meaningless textures. 291 Whereas our Experiment 2 showed that the former can survive large differences between the 292 293 orientation contents of adaptor and test, the latter cannot [19].

294

Our results are unique in the literature on the appearance of uprightness, because they show 295 296 that the global orientation of a scene can be encoded separately from its local feature content. It is assumed that information about scene orientation is embedded in the early global percept 297 of scene layout, a property which is rapidly extracted when looking at a scene [17, 28]. Based 298 on this assumption, at present, we can only speculate regarding where selectivity for the 299 orientation of natural scenes arises in the brain. One possible candidate is the 300 Parahippocampal Place Area, which is thought to encode scene layout rather than object 301 content [29]. In support of this, such scene selective regions are known to be responding 302 similarly to scenes containing only close-to-vertical or close-to-horizontal orientations [30], 303 304 akin to the stimuli we used here. Different local feature content can therefore lead to the encoding of similar global spatial layout in scenes, which presumably is what led to a 305 repulsive TAE from vertically filtered adaptors on horizontally filtered tests. 306

307

308 As noted in the introduction, the TAE is routinely invoked as a manifestation of the mutual 309 inhibition between visual mechanisms selective for orientation. Consequently, the natural

310 conclusion to draw from our results is that there must be mechanisms selective for the orientations of images with meaningful, higher-order structure. Of course, we cannot say 311 whether those mechanisms are mutually inhibitory, or whether the TAE for natural scenes 312 should be attributed to their modulation of lower-level mechanisms. Indeed, other authors 313 have invoked pre-saccadic remapping in space [18], top-down modulation of low-level 314 feature detectors through feedback from form processing regions [19] and selective attention 315 [26] in attempts to explain how the TAE can survive the spatial separation of adaptor and 316 317 tests.

318

One further possibility is normalization. Extensive real-world experience with close-toupright scenes (canonical orientation) may have resulted in the establishment of uprightness as a norm against which other orientations are compared. Exposure to tilted scenes may simply shift the subjective norm of uprightness towards the tilted direction, which then results in an objectively upright scene seen as tilted away. Indeed, Asch and Witkin [15] report that tilted scenes eventually appear upright over extended viewing, implying normalizing towards uprightness.

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Data accessibility:
Raw behavioural data from both experiments in the study can be accessed at datadryad.org:
doi:10.5061/dryad.h630v
Competing interests:
We declare that we have no competing interests.
Authors' Contributions:
All four authors developed the concept and contributed to the study design. AMHI performed
data collection. AMHI, JAS and IM were involved in the analysis and interpretation of
results. AMHI wrote the manuscript and JAS, MH and IM edited it. All authors gave final
approval for publication.
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Table 1. Group level statistics for repuision in Experiment 1 and 2										
		Repulsion (<i>R</i>)								
	Condition	Ν	mean R (°)	t - statistic (R > 0)	<i>p</i> - value	Cohen's d	paired <i>t</i> - statistic	<i>p</i> - value	Cohen's d	
nt 1	Same house	7	1.13	2.25	0.066*	0.85				
Experiment 1	Different house	6	1.31	3.62	0.015	1.48				
Exp	Different SF house	6	1.31	4.90	0.004	2.00				
int 2	Orthogonal house	10	0.65	4.11	0.003	1.30	2.42	0.039	1.16	
Experiment 2	Phase- scrambled house	10	0.20	2.68	0.025	0.85				

453 Table 1. Group level statistics for repulsion in Experiment 1 and 2

454 Notes: *N* denotes the number of observers in each condition. The asterisk (*) denotes that the 455 *p* value was approaching significance. Removing observer IM from analysis makes the p =

456

0.002.

457

458 Table 2. Group level statistics for conspicuousness in Experiment 1 and 2

		Conspicuousness (CI)								
	Condition	Ν	mean <i>CI</i> (JND)	<i>t</i> - statistic (<i>CI</i> > 0)	<i>p</i> - value	Cohen's	paired t statistic	<i>p</i> - value	Cohen's d	
nt 1	Same house	7	0.26	2.42	0.052*	0.91				
Experiment 1	Different house	6	0.27	4.24	0.008	1.73				
Exp	Different SF house	6	0.33	5.84	0.002	2.38				
ent 2	Orthogonal house	10	0.21	4.36	0.002	1.38	2.88	0.018	1.30	
Experiment 2	Phase- scrambled house	10	0.06	2.45	0.037	0.77				

⁴⁵⁹ Notes: *N* denotes the number of observers in each condition. The asterisk (*) denotes that the 460 *p* value was approaching significance. Removing observer IM from analysis makes the p =461 0.003.

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463

465 **Table and figure captions**

466

467 Table 1. Group level statistics for repulsion in Experiment 1 and 2

468

Table 2. Group level statistics for conspicuousness in Experiment 1 and 2

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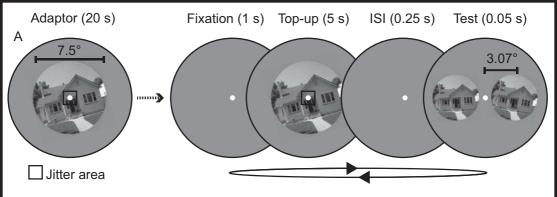
471 Figure 1. (A) Stimulus configuration and timeline of a sample trial from Experiment 1. (B)472 Five different house scenes used across the different conditions in the study.

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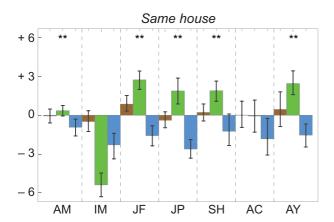
Figure 2. (A) Maximum likelihood estimates of perceptual bias for baseline (brown), CW (green) and CCW (blue) blocks from the 3 conditions in Experiment 1. Error bars are bootstrapped 95% confidence intervals. Single asterisks (*) denote after-effects significant at the $\alpha = 0.05$ level for repulsion. Double asterisks (**) denote after-effects also significant at the $\alpha = 0.001$ level for repulsion. (B) Examples of adaptors and test stimuli used in each of the conditions tested (where necessary, contrast has been amplified for visibility).

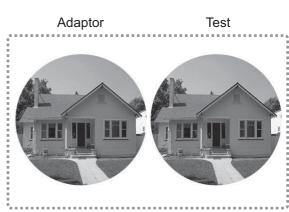
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Figure 3. Maximum likelihood estimates of perceptual bias for baseline (brown), CW (green) and CCW (blue) blocks from (A) the *orthogonal house* and (B) the *phase-scrambled house* conditions in Experiment 2. Error bars are bootstrapped 95% confidence intervals. Single asterisks (*) denote after-effects significant at the $\alpha = 0.05$ level for repulsion. Double asterisks (**) denote after-effects also significant at the $\alpha = 0.001$ level for repulsion. Examples of CW-tilted adaptors with untilted test stimuli used in each condition are illustrated to the right. The image number used for each observer is given below their initials.

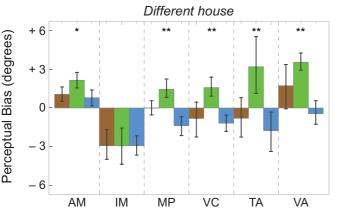




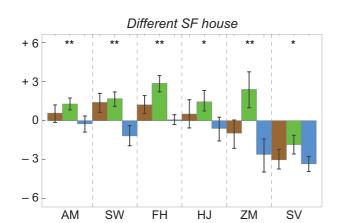




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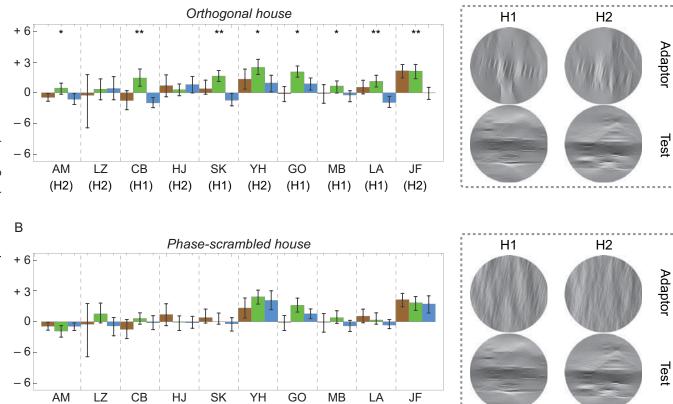












(H1)

(H1)

(H2)

Perceptual Bias (degrees)

(H2)

(H2)

(H1)

(H2)

(H1)

(H2)

(H1)