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## The Mini Alcohol Craving Experience questionnaire: Development and clinical

3

## application

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5 Jason M. Coates<sup>1,2</sup>, Matthew J. Gullo<sup>1,3</sup>, Gerald F.X. Feeney<sup>1,3</sup>, David J. Kavanagh<sup>4</sup>, Ross  
6 McD. Young<sup>3,5</sup>, Genevieve A. Dingle<sup>1,2</sup>, Jon May<sup>6</sup>, Jackie Andrade<sup>6</sup>, Dixie J. Statham<sup>7</sup>, &  
7 Jason P. Connor<sup>1,3,8</sup>

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9 <sup>1</sup> Centre for Youth Substance Abuse Research, The University of Queensland, Brisbane,  
10 Australia

11 <sup>2</sup> School of Psychology, The University of Queensland, Brisbane, Australia

12 <sup>3</sup> Alcohol and Drug Assessment Unit, Department of Medicine, Princess Alexandra Hospital,  
13 Brisbane, Australia

14 <sup>4</sup> Centre for Children's Health Research, Institute of Health & Biomedical Innovation and  
15 School of Psychology & Counselling, Queensland University of Technology, Brisbane,  
16 Australia

17 <sup>5</sup> Faculty of Health, Queensland University of Technology, Brisbane, Australia

18 <sup>6</sup> School of Psychology, Cognition Institute, Plymouth University, Plymouth, UK

19 <sup>7</sup> School of Social Sciences, University of the Sunshine Coast, Sunshine Coast, Australia

20 <sup>8</sup> School of Medicine, The University of Queensland, Brisbane, Australia

21 Corresponding Author:

22 Professor Jason Connor

23 Email: [jason.connor@uq.edu.au](mailto:jason.connor@uq.edu.au):

24 Centre for Youth Substance Abuse Research, The University of Queensland, K-Floor Mental  
25 Health Centre, Royal Brisbane and Women's Hospital, Brisbane, Australia QLD 4029

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## ABSTRACT

27           **Background:** Standardised alcohol craving scales are rarely used outside of research  
28 environments despite recognised clinical utility. Scale length is a key barrier to more  
29 widespread application. A brief measure of alcohol craving is needed to improve research and  
30 treatment of Alcohol Use Disorders (AUDs). Grounded in the Elaborated Intrusion Theory of  
31 Desire, the Alcohol Craving Experience (ACE) questionnaire comprises two 11-item self-  
32 report scales which assess past-week frequency and maximum strength of alcohol craving.  
33 This study aimed to create a brief version of the ACE while maintaining psychometric  
34 integrity and clinical utility.

35           **Methods:** Patients attending a university hospital alcohol and drug out-patient service  
36 for treatment of AUD completed the ACE as part of a questionnaire battery. Three patient  
37 samples were utilised: 519 patients with pre-treatment and outcome data; 228 patients with  
38 pre-treatment data; and 66 patients who completed the ACE at treatment sessions one and  
39 two.

40           **Results:** The Frequency scale of the ACE possessed greater clinical utility and  
41 predictive validity than the Strength scale. Revision of the Frequency measure produced a 5-  
42 item 'Mini Alcohol Craving Experience' (MACE) questionnaire. Satisfactory validity  
43 (construct, predictive, concurrent, convergent, and incremental) and reliability (internal and  
44 test-retest) was maintained. A one standard deviation increase in pre-treatment MACE score  
45 was associated with a 54 percent increase in the odds of patient lapse or dropout.

46           **Conclusions:** The MACE provides a brief, theoretically and psychometrically robust  
47 measure of alcohol craving suitable for use with AUD populations in time-limited clinical  
48 and research settings.

49           **Keywords:** Alcohol Use Disorder, Craving, Urge, Measurement, Scale development

## INTRODUCTION

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Craving is a robust marker of substance dependence severity and is implicated in treatment relapse (Flannery et al. 2003; Law et al. 2016; Yoshimura et al. 2016). The Diagnostic and Statistical Manual of Mental Disorders, Fifth Edition (DSM-5) recently included ‘craving, or a strong desire or urge to use a substance’ as a diagnostic criterion for Substance Use Disorders (American Psychiatric Association, 2013). Craving was defined as a strong desire to consume a substance that makes it difficult to think of anything else (American Psychiatric Association, 2013; Hasin et al. 2013). Craving interventions feature prominently in psychological treatments, and pharmacotherapies have been developed to target specific craving neuromechanisms (Addolorato et al. 2005; Haass-Koffler et al. 2014). After decades of experimental, clinical, and epidemiological research, accurate measurement of substance craving remains a research priority (Tiffany and Wray 2012; Kavanagh et al. 2013). Historically, craving has been measured by conceptually weak and often unstandardised methods, limiting generalisability and clinical utility (Sayette et al. 2000; Pavlick et al. 2009; Kavanagh et al. 2013). Some standardised scales have been introduced, although uptake within clinical settings has been poor (Pavlick et al. 2009; Tiffany and Wray 2012).

A national survey of U.S. addiction services found 99% considered craving in treatment planning, yet only 5% employed standardised self-report craving measures (Pavlick et al. 2009). The majority opted for single-item or non-standard open ended questions, despite well documented limitations to the reliability of these approaches (Cortina 1993; Hruschka et al. 2004). This may reflect the psychometric and theoretical weaknesses in self-report craving scales (Sayette et al. 2000; Kavanagh et al. 2013) and time burden imposed by scale administration and analysis in busy clinical environments. Alcohol Use Disorders (AUDs) are among the most prevalent Substance Use Disorders, placing a substantial burden upon global

75 mortality and disease (Connor and Hall 2015; Gowing et al. 2015; Connor et al. 2016). A  
76 brief, psychometrically sound measure of alcohol craving is needed to improve assessment,  
77 diagnosis, and treatment of AUDs.

78 Measures vary considerably in their definition of craving. In a recent review of  
79 alcohol craving scales, based on 47 papers published between 1990 and 2012, we argued that  
80 the majority contain constructs extraneous to widely applied diagnostic definitions of craving  
81 (e.g. DSM-5, ICD-10; Kavanagh 2013). These often include items measuring allied  
82 constructs, such as expectancies, intentions, and refusal self-efficacy (Kavanagh et al. 2013).  
83 Though such constructs are important within models of substance use and craving, the  
84 presence of these allied phenomena may influence accurate diagnosis of AUD and bias  
85 conclusions drawn from subsequent research. For example, the inclusion of items assessing  
86 self-efficacy (Bandura 1977) may artificially inflate the predictive utility of a scale, as self-  
87 efficacy about drinking control reliably predicts drinking behaviour (Connor et al. 2007).

88 The presence of allied addiction constructs does not necessarily compromise the  
89 validity of a craving scale. If the outcomes are interpreted in the context of a prescribed  
90 definition or with regard to a theoretical model then construct validity may be maintained.  
91 However, craving scales infrequently report a definition to which they adhere and are often  
92 developed atheoretically (Flannery et al. 1999; Rojewski et al. 2015; McHugh et al. 2016).  
93 We developed the Alcohol Craving Experience (ACE) Questionnaire to be consistent with  
94 common definitions of craving while adhering to a specified theory (Statham et al. 2011).  
95 However, administration of the 22-item ACE is likely to be too time consuming for practical  
96 use. It is proposed that reduction of the ACE would result in a theoretically and  
97 psychometrically sound measure of craving which may be easily integrated in time-limited  
98 environments.

99

100 Reflecting the Elaborated Intrusion (EI) Theory of Desire (Kavanagh et al. 2005; May  
101 et al. 2014b), the ACE measures three aspects of craving: the intensity of the drive to drink  
102 (Intensity), the presence of associated imagery (Imagery), and intrusiveness of desire  
103 cognitions (Intrusion; Statham et al. 2011). EI theory defines craving as an affectively laden  
104 cognitive event, where an object or activity and its associated pleasure or relief is in focal  
105 attention (Kavanagh et al. 2005). Consistent with neurobiological models of craving,  
106 addictive substances are believed to recruit the same physiological mechanisms that drive  
107 appetitive behaviours required for survival (Robinson and Berridge 1993). EI theory proposes  
108 that biological, environmental, and affective cues trigger intrusive desire-related cognitions  
109 which occupy attention and prompt elaboration. The subsequent elaboration process—in  
110 particular imagery—provides momentary pleasure or relief of physical and emotional  
111 discomfort (Connor et al. 2014). However, pleasure or relief from elaborative cognitions  
112 quickly dissipates. Instead, awareness is drawn to any emotional or physical deprivation and  
113 to potential actions to acquire the target. Further elaboration and intensification of the desire  
114 ensues, unless the target is acquired or attention is captured elsewhere.

115 EI theory aligns with treatment approaches such as motivational enhancement,  
116 mindfulness, acceptance-based therapies, and retraining attentional biases (Witkiewitz et al.  
117 2013; May et al. 2014b; Witkiewitz et al. 2014). Recent research has directly employed EI  
118 theory in the development of promising new craving management strategies and novel  
119 treatment approaches (Kemps and Tiggemann 2007; Knäuper et al. 2011; Kemps and  
120 Tiggemann 2013; Hsu et al. 2014; Skorka-Brown et al. 2014; Littel et al. 2016). These  
121 approaches employ non-substance imagery and sensory tasks designed to compete with  
122 craving-based imagery within the limited capacity of working memory. The information  
123 provided by the ACE may facilitate more detailed formulation, treatment planning, and  
124 monitoring of craving.



149 Cognitive Behaviour Therapy (CBT) conducted over 12 weeks. Treatment may be  
150 supplemented by pharmacotherapy (naltrexone, acamprosate, or both). The assessment  
151 battery is completed in a separate consultation prior to the first treatment session and again at  
152 the completion of treatment. All patients were over 18 years of age and met DSM-IV  
153 (American Psychiatric Association 2000) criteria for alcohol dependence. Human ethics  
154 approval was obtained (2008/125, HREC/12/QPAH/022 HREC/14/QPAH/664) and  
155 participants provided informed written consent. Sample characteristics are presented in Table  
156 1.

157

158 *Scale Reduction Sample.* This sample comprised 519 alcohol dependent patients  
159 (Table 1). All patients were over 18 years of age and met DSM-IV(American Psychiatric  
160 Association 2000) criteria for alcohol dependence. These data have been used previously in  
161 the original development of the ACE (Statham et al. 2011) and in examining craving as a  
162 mediator of change (Law et al. 2016), but have not been used to directly predict treatment  
163 outcome.

164

165 *Validation Sample.* The validation sample comprised pre-treatment data from 228  
166 consecutively treated alcohol dependent patients (Table 1). These data were employed to  
167 assess the factor structure of the ACE scales and cross-sectional relationships between  
168 variables.

169

170 *Test-Retest (TRT) Sample.* The ACE-F was administered to 66 patients at treatment  
171 sessions one and two, in-order to assess test-retest reliability of the ACE-F. Mean time  
172 between sessions was 8.40 days ( $SD = 2.86$ ).

173

174 **Insert Table 1**

175

176

177 *Measures*

178 *The Alcohol Craving Experience (ACE) questionnaire.* The ACE comprises two 11-  
179 item scales that assess the frequency (ACE-F) and peak strength (ACE-S) of desire-related  
180 cognitions over the previous week. Items load onto three classes of cognition, ‘Intensity’  
181 (items 1-3), ‘Imagery’ (items 4-8), and ‘Intrusion’ (items 9-11). Participants respond via an  
182 11-point visual analogue scale with anchors 0 (*not at all*) and 10 (*constantly/extremely*). The  
183 ACE-F and ACE-S have good internal reliability and concurrent validity, and can  
184 discriminate between problem and non-problem drinkers (Statham et al. 2011).

185

186 *The Obsessive Compulsive Drinking Scale (OCDS).* The OCDS is a 14-item self-  
187 report measure intended to reflect drinking-related obsessive and compulsive craving and  
188 behaviour (Anton et al. 1995). The OCDS has received extensive research attention and is  
189 currently the most widely used measure of alcohol craving. The OCDS has acceptable test-  
190 retest reliability, internal reliability, and concurrent validity (Anton et al. 1995; Kranzler et al.  
191 1999; Roberts et al. 1999). The OCDS cannot be considered a ‘pure’ measure of craving as  
192 extraneous constructs such as consumption, effort to resist drinking, functional interference  
193 and distress from drinking, as well as perceived control of drinking are all assessed within the  
194 scale. The first six items, comprising the Obsessions Subscale are most consistent with the  
195 clinical definitions of craving. OCDS-Obsessions is intended to assess drinking obsession  
196 related cognitions, for example, “How much of your time when you’re not drinking is  
197 occupied by ideas, thoughts, impulses, or images related to drinking?”. While less  
198 confounded than the full OCDS, OCDS-Obsessions does contain extraneous phenomena,  
199 assessing functional interference and distress caused by obsessive cognitions. OCDS-



200 Obsessions has been demonstrated to improve prediction of drinking behaviour (Flannery et  
201 al. 2003) and likelihood of relapse post treatment (Soyka et al. 2010). As OCDS-Obsessions  
202 is a widely used measure of craving and considered among the better performing craving  
203 scales (Kavanagh et al. 2013) it was employed as a concurrent measure of alcohol craving.

204

205 *The Alcohol Use Disorders Identification Test (AUDIT)*. The AUDIT is a 10-item,  
206 self-report measure assessing recent alcohol use, symptoms of alcohol dependence, and  
207 alcohol related problems (Saunders et al. 1993). The AUDIT has sound internal reliability,  
208 sensitivity and specificity, and discriminant validity (Saunders et al. 1993). Higher scores  
209 indicate increased risk of harmful or hazardous drinking.

210

211 *The Beck Depression Inventory - Second Edition (BDI-II)*. The BDI-II is a 21-item  
212 self-report measure assessing attitudes and behaviours symptomatic of depression (Beck et al.  
213 1996). The BDI-II is a well validated measure demonstrating strong test-retest and internal  
214 reliability, as well as good concurrent, content, discriminant, and construct validity (Beck et  
215 al. 1988; Beck et al. 1996).

216

217 *The State Anxiety Scale (S-Anxiety)*. The S-Anxiety Scale of the State Trait Anxiety  
218 Inventory (STAI) comprises 20 self-report items assessing the respondent's current state of  
219 anxiety (Spielberger 1983). The S-Anxiety has acceptable internal and test-retest reliability,  
220 as well as content, discriminant, and construct validity (Spielberger 1983; Oei et al. 1990;  
221 Barnes et al. 2002).

222

223 *Procedure*

224           *Scale Reduction.* To best maintain consistency of the measured construct, an initial  
225 step involved selection of a form of the ACE for further refinement (ACE-F or ACE-S). Each  
226 form was evaluated based on perceived clinical utility and predictive validity. Decisions  
227 guiding subsequent item reduction were informed by the following rationale: (a) to enhance  
228 construct validity, items with the greatest face validity and theoretical importance within EI  
229 theory were prioritised; (b) to maximise the sensitivity and clinical utility of a reduced scale,  
230 the most highly endorsed items were also prioritised for retention; (c) to enhance predictive  
231 validity, the capacity of items to discriminate between patients who lapsed or withdrew from  
232 treatment and those who were abstinent throughout treatment was also considered. Data  
233 analyses within this step utilised the Scale Reduction Sample.

234

235           *Scale Evaluation.* Reduced models were further evaluated based on construct,  
236 predictive, concurrent, and convergent validity, as well as internal and test-retest reliability.  
237 Predictive validity of OCDS-Obsessions was also assessed for concurrent comparison. Data  
238 analysis within this step utilised the Validation and Test-Retest samples.

239

240           *Scale Selection.* The shortest scale maintaining psychometric integrity would be  
241 selected as the final reduced version.

242

### 243 *Data Analysis*

244           Analyses were conducted in SPSS version 22. Confirmatory factor analyses (CFA)  
245 were conducted in R version 3.2.1 (R Core Team 2015), package extension *lavaan* .5-18  
246 (Rosseel 2012). As the distributions of all ACE item and scale scores were significantly  
247 negatively skewed, statistical procedures robust to non-normal distributions were utilised.  
248 CFA Models were compared using changes in  $\chi^2$  /df ratios (smaller values indicating

249 improved fit; Carmines and McIver 1981), Comparative Fit Indices (CFI, values >.93  
 250 indicating good fit; Hu and Bentler 1999) , Standardised Root Mean Square Residual  
 251 (SRMR; Values <.07 indicating good fit; Hu and Bentler 1999), Root Mean Square Error of  
 252 Approximation (RMSEA; values <.07 indicating good fit; Hu and Bentler 1999), and Akaike  
 253 Information Criterion (AIC; smaller values indicating improved fit; Bozdogan 1987).

254

## 255 RESULTS

### 256 *Scale Reduction*

257 *Subscale-Selection.* As the ACE-S asks the respondent to report on only the most  
 258 severe episode of past week craving, it is influenced by contextual factors such as situational  
 259 cues and novel stressors. Clinical value of this method is drawn from the isolation of a  
 260 specific time-period where the patient may be most vulnerable to lapse. Alternatively, the  
 261 ACE-F assesses the perceived frequency of craving symptoms over the past week, providing  
 262 a more general overview of the patients craving experience. The ACE-F was subsequently  
 263 identified as the preferred scale for reduction, based on its perceived benefit as a measure  
 264 more sensitive to change in the patient's typical craving experience.

265 Using the Scale Reduction Sample, separate logistic regression analyses were  
 266 employed to assess the capacity of pre-treatment ACE scale scores to predict the likelihood of  
 267 treatment lapse relative to patients who were abstinent throughout treatment. Patients who  
 268 discontinued treatment without record of lapse were conservatively included within the lapse  
 269 group. All scale scores were standardised to facilitate the comparison of effects. AUDIT  
 270 scores and medication status were included as covariates, but did not significantly improve  
 271 upon the intercepts-only model ( $\chi^2(2) = 0.26, p = .877, Nagelkerke R^2 = .001$ ; Table 2,  
 272 Baseline Model). Inclusion of either the ACE-S ( $\Delta\chi^2(1) = 18.71, \Delta p = <.001, Nagelkerke$   
 273  $\Delta R^2 = .054$ , Table 2, Model 1) or ACE-F ( $\Delta\chi^2(1) = 21.68, \Delta p = <.001, Nagelkerke \Delta R^2 =$

274 .062, Table 2, Model 2) significantly improved the predictive power of the model. As Model  
275 2 appeared to explain more variance than Model 1, the ACE-F was added to Model 1 in an  
276 additional step to examine if it would account for significantly more variance than the ACE-  
277 S. The addition of the ACE-F to Model 1, saw the ACE-F become the dominant predictor  
278 within the model, though predictive power was not significantly improved ( $\Delta\chi^2(1) = 3.63$ ,  $\Delta p$   
279  $= .057$ , Nagelkerke  $\Delta R^2 = .011$ , Table 2, Model 3). The ACE-F was subsequently selected for  
280 further refinement.

281

282

### Insert Table 2

283

284 *Item Importance.* Prior to item reduction, the structure and items central to the  
285 theoretical foundation of the scale were considered. At least one item from each sub-scale  
286 was retained to represent each factor. Items 3 and 9 (Table 3) were prioritized for retention  
287 due to high semantic consistency to the Intensity and Intrusion factors respectively. Multiple  
288 items of the Imagery factor would be retained to capture potential individual differences in  
289 the most prevalent imagery modalities involved in alcohol craving.

290

291 *Feature Prevalence.* Medians and interquartile ranges for all ACE-F items are  
292 presented in Table S1 within the online supplementary material. While all items had an  
293 interquartile range of at least 4 on the 11-point scale, most also received a large proportion of  
294 'not at all' responses. To identify which items were most representative of common craving  
295 symptoms among patients with AUD, the endorsement rates (ERs; proportion of non-zero  
296 responses to each item) were also calculated. McNemar's  $\chi^2$  was utilised to identify  
297 significant differences between items in the prevalence of endorsement rates within each  
298 factor. Within the Intensity factor, the endorsement rate of Item 2 (80.2%) was significantly

299 lower than Item 3 (86.1%,  $p < .001$ ), while Items 1 (87.6%) and 3 could not be distinguished  
300 ( $p = .169$ ). Comparisons of endorsement rates of items within the Imagery factor revealed all  
301 were significantly different ( $p < .001$ ), with the exception of the most highly endorsed, items  
302 4 (80.9%) and 8 (80.1%,  $p = .716$ ). Within the Intrusion factor, item 11 was the least  
303 endorsed factor (75.8%,  $p < .001$ ) while items 9 (84.9%) and 10 (83.8%) could not be  
304 differentiated ( $p = .291$ ).

305 Separate Mann-Whitney  $U$  tests revealed that the mean rank of patients who lapsed or  
306 withdrew from treatment was significantly higher for every item than those who completed  
307 treatment abstinent (Table 3). Steiger's  $Z$  revealed no significant differences in the size of the  
308 effects between items.

309

310

### Insert Table 3

311

312 *Item Reduction.* To maximize sensitivity of the reduced craving measure items with  
313 the highest endorsement rates were given greater priority for retention to minimise the  
314 number of 'not at all' responses within the reduced scale. Based on feature prevalence and  
315 consistency with the overarching factors, items 3 and 9 were retained to represent the  
316 Intensity and Intrusion factors respectively. The three imagery items with the highest  
317 endorsement rates (4, 5, and 8) were retained to comprise the initial Imagery factor.

318 A sequential logistic regression was employed to assess the capacity for the selected  
319 items to predict alcohol lapse in the Scale Reduction Sample. Addition of the items intended  
320 to comprise the reduced ACE (items: 3, 4, 5, 8, 9) to the Baseline Model (Table S2)  
321 significantly improved predictive power of the model ( $\Delta\chi^2(5) = 21.49$ ,  $\Delta p < .001$ ,  
322 *Nagelkerke*  $\Delta R^2 = .061$ , Model 4, Table S2). To assess whether the model could be improved  
323 with the inclusion of additional ACE items, the remaining items were included using forward

324 entry. Sequential inclusion of items 1 ( $\Delta\chi^2(1) = 7.61, \Delta p = .006, \text{Nagelkerke } \Delta R^2 = .023,$   
325 Model 5, Table S2) and 10 ( $\Delta\chi^2(1) = 9.84, \Delta p = .002, \text{Nagelkerke } \Delta R^2 = .027,$  Model 6, Table  
326 S2) would significantly improve the final model ( $\chi^2(9) = 39.20, p < .001, \text{Nagelkerke } R^2 =$   
327 .111).

328

### 329 *Scale Evaluation*

330 *Validity.* To assess the construct validity of the initial five-item scale, the seven-item  
331 scale, and the complete ACE-F, confirmatory factor analyses were performed utilising the  
332 Validation Sample. Maximum likelihood estimation with robust standard errors and a  
333 Satorra-Bentler scaled test statistic were employed to reduce the effects of non-normality.  
334 Model fit statistics are presented in table 4, and parameter estimates are summarised in the  
335 supplementary material. For the 11 and 7 item scales, the three-factor solution provided a  
336 better fit to the data than a unifactorial model (Table 4). For the five item scale, both  
337 solutions showed comparable fit. The CFI, RMSEA, SRMR, and AIC fit statistics all  
338 improved through reduction. No covariance between error terms was specified in any of the  
339 models. These results support previous studies validating the three-factor structure of the  
340 ACE (Statham et al. 2011; May et al. 2014a), though when reduced to a five-item scale, it  
341 could equally reflect a global construct of craving within a single factor (Figure 1).

342

343 **Insert Table 4**

344 **Insert Figure 1**

345

346 Data from the Validation Sample indicated that all scales had significant ( $p < 0.001$ )  
347 large positive correlations with OCDS-Obsessions, indicating an acceptable level of  
348 concurrent validity ( $r = 0.60$  to  $0.58$ ). Convergent validity was demonstrated by significant ( $p$

349 < 0.01) small to moderate positive correlations with the AUDIT ( $r = 0.22$  to  $0.20$ ) and  
 350 significant ( $p < 0.001$ ) moderate correlations with measures of anxiety (S-Anxiety:  $r = 0.40$   
 351 to  $0.38$ ) and depression (BDI:  $r = 0.39$  to  $0.38$ ). The strength of the correlations did not  
 352 significantly differ between the three ACE versions (Steiger's  $Z$ ,  $p < .05$ ), indicating that  
 353 convergent and concurrent validity of the ACE was not significantly affected by scale  
 354 reduction.

355 Utilising the Scale Reduction Sample predictive validity of the scales administered  
 356 pre-treatment was assessed by logistic regressions with the outcomes 'complete treatment  
 357 abstinent' and 'lapsed or discontinued treatment'. When independently added to the Baseline  
 358 Model, the five-item ( $\Delta\chi^2(1) = 15.17$ ,  $\Delta p < .001$ , *Nagelkerke*  $\Delta R^2 = .044$ , Model 7, Table 5),  
 359 seven-item ( $\Delta\chi^2(1) = 20.19$ ,  $\Delta p < .001$ , *Nagelkerke*  $\Delta R^2 = .058$ , Model 8, Table 5), and 11-  
 360 item (Model 2, Table 2) scales all significantly improved predictive power of the model.  
 361 Predictive power of OCDS-Obsessions was also assessed for concurrent comparison.  
 362 Addition of OCDS-Obsessions significantly improved upon the Baseline Model ( $\Delta\chi^2(1) =$   
 363  $7.78$ ,  $\Delta p = .005$ , *Nagelkerke*  $\Delta R^2 = .022$ , Model 9, Table 5). The incremental validity of each  
 364 scale was assessed by systematically adding the weaker of two scales, based on *Nagelkerke's*  
 365  $R^2$ , to the Baseline Model, followed by the next strongest scale in step two. The 5-item ACE-  
 366 F was demonstrated to significantly improve upon the predictive power of OCDS-Obsessions  
 367 ( $\Delta\chi^2(1) = 7.35$ ,  $\Delta p = .007$ , *Nagelkerke*  $\Delta R^2 = .044$ , Model 10, Table 5) and the 7-item scale  
 368 significantly improved upon the 5-item ( $\Delta\chi^2(1) = 15.43$ ,  $\Delta p < .001$ , *Nagelkerke*  $\Delta R^2 = .088$ ,  
 369 Model 11, Table 5). The 11-item scale did not improve upon the seven-item scale ( $\Delta\chi^2(1) =$   
 370  $1.19$ ,  $\Delta p = .173$ , *Nagelkerke*  $\Delta R^2 = .064$ , Model 12, Table 5).

371

372

**Insert Table 5**

373

374           *Reliability.* Internal consistency was assessed using the Validation Sample.  
375 Cronbach's Alpha was above .90 for all scales with only minor reductions in the reduced  
376 scales ( $\alpha = 0.95$  to  $0.92$ ). Test-Retest reliability utilised session one and two data from 66  
377 patients. Correlations between session one and session two ACE scores indicated that test-  
378 retest reliability was acceptable across all scales ( $r = 0.731$  to  $0.725$ ). Steiger's  $Z$  revealed no  
379 significant changes in scale test-retest reliability following reduction.

380

### 381 *Scale Selection*

382           The procedures conducted indicate that the ACE-F may be reduced to as few as five  
383 items while maintaining theoretical and psychometric integrity. The five-item scale, termed  
384 the Mini Alcohol Craving Experience (MACE), was chosen as the most suitable short-form  
385 scale for assessment of craving in AUD populations.

386

387

## DISCUSSION

388           In place of the two 11-item forms of the ACE, a brief five-item measure of craving  
389 was validated (MACE). The MACE maintained high construct, predictive, concurrent, and  
390 convergent validity. High internal and test-retest reliability consistent with the ACE-F was  
391 also demonstrated. The MACE measures the frequency of past week craving including  
392 intense urges, imagery, and intrusiveness of craving related cognitions (Kavanagh et al.  
393 2005). The MACE is simple to administer and may be completed in less than 60 seconds,  
394 reducing time burden on respondents, health professionals, and researchers.

395           In addition to its brevity, the MACE maintains several strengths uncommon among  
396 current craving instruments, including a strong theoretical model and absence of drinking  
397 constructs known to confound craving measurement (Sayette et al. 2000; Kavanagh et al.  
398 2013). By retaining the items most representative of the ACE factors, and monitoring the



399 resultant model fit, the MACE preserved the construct validity of the ACE. The MACE  
400 subsequently retains the capacity for unique insight into intensity and intrusiveness of patient  
401 craving, as well and key elements of craving based imagery. This information may inform  
402 case formulation and treatment planning.

403 Predictive validity is infrequently examined in existing craving measures. Higher  
404 scores on the MACE were predictive of increased risk of lapse or dropout from treatment in  
405 this alcohol dependent sample. A one standard deviation increase in MACE score was  
406 associated with a 54% increase in the odds of lapse or discontinuation of treatment; relative  
407 to OCDS-Obsessions, where a one standard deviation score increase was associated with a  
408 10% increase in risk. The practical interpretation of this result is that for every one-point  
409 increase on the MACE pre-treatment (maximum score = 50), the odds of a patient completing  
410 treatment abstinent reduced by 3.1 percent. The MACE may therefore assist addiction  
411 professionals to better assess risk of relapse in their patients.

412 Few craving measures assess test-retest reliability. The MACE deliberately measures  
413 past week frequency of craving, under the assumption that this will have greater stability and  
414 subsequently be a more reliable indicator of change than single time point assessments. The  
415 correlation of session one and two MACE scores was  $r = 0.73$ , and is interpreted as an  
416 acceptable degree of stability within the clinical context. Given the prominence of craving  
417 within clinical and research settings, a measure of craving sensitive to change over time is  
418 greatly needed. The MACE may enhance the validity of studies assessing the efficacy of  
419 craving interventions, and improve monitoring of patients' treatment response in clinical  
420 settings.

421 As this study was conducted in a hospital outpatient clinic, the samples provided  
422 optimal, clinically relevant data. However, the practical nature of the research design  
423 introduced some limitations. The samples predominantly comprised middle-aged men with

424 poor social or occupational functioning and moderate to severe alcohol dependence. Future  
425 studies should investigate the MACE in more diverse patient populations, as craving profiles  
426 may vary across problem severity, age, culture, social-occupational status. An additional  
427 limitation is that follow up data of patients who dropped out were not available, and were  
428 conservatively recorded as having lapsed. Assessment of test-retest reliability was also  
429 impaired by the treatment setting. An increased focus on drinking and attempts to change  
430 drinking behaviours is likely to have increased variance in patient craving from session one to  
431 two. While this is hypothesised to have led to the underestimation of the MACE's stability  
432 future research should assess participants under stable conditions with tightly controlled time  
433 points. Further research is also needed to examine the performance of the MACE as a stand-  
434 alone measure. As the MACE was only assessed as a sub-selection of the full ACE, the extent  
435 to which the variance of the retained items is influenced by the excluded items is unknown.  
436 Finally, while craving frequency presents ongoing challenges to the control of drinking, very  
437 intense peak levels also constitute significant risk. Utilising both frequency and strength  
438 forms of the ACE is recommended when time permits, as they offer a more comprehensive  
439 assessment of the patient's experience of craving. The MACE and ACE scales, scoring  
440 instructions, and normative data are included in the online supplementary material.

441 A final recommendation, which applies to the use of all craving measures, is that scale  
442 administrators, researchers and clinicians alike, carefully interpret scale scores in light of the  
443 definition and theory under which they are proposed. It is argued that unclear definitions, and  
444 the absence of theoretical models have impaired craving measurement to date, confounding  
445 the craving construct as it is widely understood (Tiffany and Wray 2012; Kavanagh et al.  
446 2013). Interpreting ACE scores in the context of the Elaborated Intrusion Theory of Desire  
447 (Kavanagh et al. 2005) will improve understanding of the proposed construct of craving and  
448 enhance its clinical utility.

449           The Mini Alcohol Craving Experience (MACE) reflects the key theoretical elements  
450 of the ACE, while maintaining the best performing items and preserving psychometric  
451 integrity. Key strengths of the MACE include excellent construct validity, predictive validity,  
452 and acceptable test-retest reliability. In conjunction with its brevity, these features make the  
453 MACE ideal for use with AUD populations in time limited clinical and research  
454 environments.

455

456

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462

463

#### CONFLICTS OF INTEREST

464           There are no conflicts of interest to declare.

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466

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- 594

595 **Table 1.** Patient sample characteristics

Sample characteristics	Scale Reduction Sample <i>n</i> = 519	Validation Sample <i>n</i> = 228	TRT Sample <i>n</i> = 66
Mean Age, years (SD)	39.82 (11.59)	44.39 (10.82)	45.48 (10.03)
Sex, female	171 (32.9%)	84 (36.8%)	22 (33.3)
Married/ <i>De-facto</i>	184 (35.5%%)	82 (36.0%)	25 (37.9%)
Education			
Degree	70 (13.5%)	47 (20.5%)	17 (25.8%)
Diploma/Certificate	52 (10.0%)	16 (7.1%)	6 (9.1%)
Senior Secondary (Year 12)	157 (30.3%)	71 (31.1%)	22 (33.3%)
Junior Secondary (Year 10)	190 (36.6%)	82 (36.0%)	17 (25.8%)
Primary (Year 7)	33 (6.4%)	11 (4.8%)	4 (6.1%)
Unemployed	103 (19.8%)	44 (19.3%)	15 (22.7%)
Mean Alcohol (grams) per drinking day (SD)	147.07 (88.90)	169.80 (100.93)	196.12 (119.71)
Median Baseline ACE-F (IQR)	39 (48.00)	42.00 (46.75)	43.50 (45.50)
Mean Baseline AUDIT (SD)	27.25 (8.6)	29.38 (7.01)	27.47 (10.28)
Mean Baseline OCDS-Obsessions (SD)	7.82 (4.47)	8.82 (4.36)	8.46 (4.76)
Medication Prescribed*	315 (60.7%)	25 (11.0%)	10 (15.2%)

596 \*The Scale Reduction Sample records medication (naltrexone/acamprosate/both) if it is prescribed at any point during treatment. Medication is  
597 only counted in the Validation and TRT samples if it was taken in the week prior to assessment. As the Validation sample assessment occurred  
598 prior to commencement of behavioural treatment and TRT sample was assessed in Session 1, the majority of patients had not yet been prescribed  
599 pharmacotherapy.

600

601 **Table 2.** Summary of hierarchical logistic regression models assessing predictive validity of the ACE-F and ACE-S.

	$\beta$ (SE)	95% CI for Odds Ratio		
		Lower	Odds Ratio	Upper
Baseline Model				
Constant	1.18*** (.13)		3.26	
Medication	0.11 (.22)	0.73	1.12	1.71
AUDIT	-0.00 (.11)	0.81	1.00	1.23
Model 1				
Constant	1.19*** (.14)		3.28	
Medication	0.23 (.22)	0.81	1.26	1.96
AUDIT	-0.04 (.11)	0.77	0.96	1.20
ACE-S	0.46*** (.11)	1.28	1.59	1.97
Model 2				
Constant	1.21*** (.14)		3.34	
Medication	0.23 (.22)	0.81	1.26	1.95
AUDIT	-0.05 (.11)	0.76	0.95	1.18
ACE-F	0.53*** (.12)	1.34	1.69	2.14
Model 3				
Constant	1.2*** (.14)		3.32	
Medication	0.24 (.23)	1.27	1.27	1.98
AUDIT	0.05 (.11)	0.95	0.95	1.18
ACE-S	0.15 (.19)	1.17	1.17	1.70
ACE-F	0.39 (.20)	1.48	1.48	2.21

Note: \*  $p < .05$ , \*\*  $p < .01$ , \*\*\*  $p < .001$ ,

604

605 **Table 3.** Mean rank comparison of abstinent patients and those who lapsed or dropped out of treatment across all ACE-F items scores.

How often did these things happen over the last week?	Complete Abstinent		Lapse or Dropout		U	Z	<i>p</i>	<i>r</i>
	<i>n</i>	Mean Rank	<i>n</i>	Mean Rank				
1. Did you want a drink?	118	196.24	398	276.96	16135.00	-5.19	<.001	-0.23
2. Did you think about needing a drink?	118	203.00	399	275.56	16933.00	-4.67	<.001	-0.20
3. Did you have an urge to drink?	118	203.95	399	275.28	17045.00	-4.58	<.001	-0.20
4. Did you picture alcohol or drinking?	118	215.42	399	271.89	18398.50	-3.64	<.001	-0.16
5. Did you imagine what it would taste like?	118	215.79	398	271.16	18442.50	-3.59	<.001	-0.16
6. Did you imagine what it would smell like?	118	217.61	399	271.24	18656.50	-3.54	<.001	-0.16
7. Did you imagine what it would feel like in your mouth or throat?	118	214.71	399	272.10	18315.00	-3.74	<.001	-0.16
8. Did you imagine how your body would feel if you had a drink?	118	223.04	398	269.01	19298.00	-2.96	0.003	-0.13
9. When you thought about alcohol over the last week, how often were the thoughts intrusive?	117	223.46	388	261.91	19241.50	-2.51	0.012	-0.11

10. When you thought about alcohol over the last week, how often were you trying not to think about alcohol?	117	211.29	398	271.73	17818	-3.88	<.001	-0.17
11. Did you find it hard to think about anything else?	118	203.59	399	275.56	17003	-4.55	<.001	-0.20

606

607 **Table 4.** Robust fit indices for the 3-factor and unifactorial structures of the ACE scales ( $n = 228$ ).

Scale	$\chi^2$ ( $df$ )	$\chi^2 / df$	$p$	CFI	RMSEA	SRMR	AIC
<b>ACE-F 11</b>							
Unifactorial	302.13 (44)	6.87	<.001	0.898	0.160	0.069	11236.7
3-Factor	158.92 (41)	3.88	<.001	0.954	0.112	0.056	11013.50
<b>ACE-F 7</b>							
Unifactorial	78.91 (14)	5.64	<.001	0.955	0.143	0.040	7321.29
3-Factor	35.59 (11)	3.24	<.001	0.983	0.099	0.027	7265.35
<b>ACE-F 5</b>							
Unifactorial	23.23 (5)	4.65	<.001	0.983	0.126	0.026	5197.70
3-Factor	23.47 (4)	5.87	<.001	0.982	0.146	0.026	5199.57

608

609

610 **Table 5.** Summary of hierarchical logistic regression models assessing predictive validity of the reduced ACE-F Scales and OBS.

611

	$\beta$ (SE)	95% CI for Odds Ratio		
		Lower	Odds Ratio	Upper
Model 7				
Constant	1.19*** (.14)		3.28	
Medication	0.22 (.22)	0.8	1.25	1.93
AUDIT	-0.04 (.11)	0.78	0.96	1.19
ACE-F-5 item	0.43*** (.12)	1.23	1.54	1.93
Model 8				
Constant	1.19*** (.14)		3.3	
Medication	0.24 (.23)	0.82	1.27	1.98
AUDIT	-0.04 (.11)	0.77	0.96	1.19
ACE-F-7 item	0.50*** (.12)	1.31	1.65	2.06
Model 9				
Constant	1.18*** (.14)		3.24	
Medication	0.20 (.22)	0.79	1.23	1.9
AUDIT	-0.07 (.11)	0.75	0.93	1.16
OBS	0.31** (.11)	1.09	1.37	1.71
Model 10				
Constant	1.19*** (.14)		3.29	
Medication	0.225 (.23)	0.8	1.25	1.95
AUDIT	-0.06 (.11)	0.76	0.95	1.18
OBS	0.10 (.14)	0.84	1.1	1.44
ACE-F-5 item	0.37** (.14)	1.11	1.45	1.9

## Model 11

Constant	1.22 (0.14)		3.38	
Medication	0.26 (0.23)	0.83	1.29	2.02
AUDIT	-0.03 (0.11)	0.77	0.97	1.21
ACE-F-5 item	-2.21 (0.7)	0.03	0.11	0.43
ACE-F-7 item	2.67 (0.7)	3.65	14.39	56.77

## Model 12

Constant	1.22 (0.14)		3.37	
Medication	0.21 (0.23)	0.8	1.24	1.93
AUDIT	-0.06 (0.11)	0.76	0.94	1.17
ACE-F-7 item	-0.40 (0.67)	0.18	0.67	2.48
ACE-F-11 item	0.93 (0.69)	0.66	2.55	9.82

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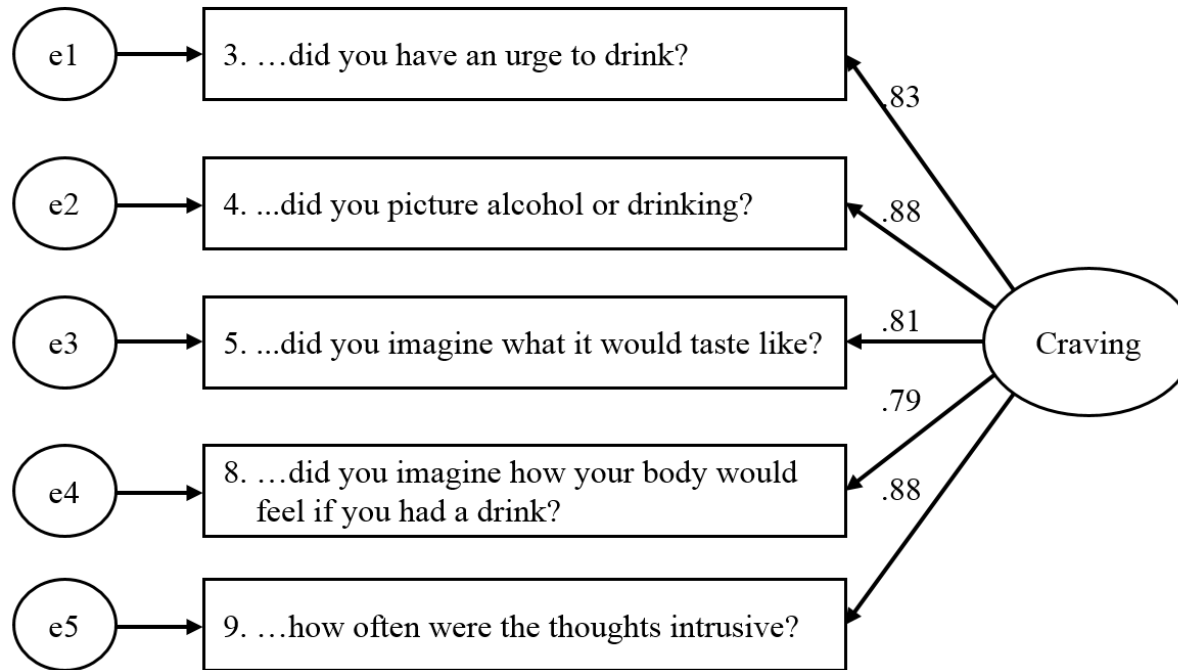
Note: \*  $p < .05$ , \*\*  $p < .01$ , \*\*\*  $p < .001$ .

612

613

614

How often did these things happen over the last week?



615

616 **Figure 1.** Unifactorial model of the 5-item ACE-F with standardised parameter. All paths are significant at  $p < .001$ .