

# The Democratic Effect of Direct Democracy

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## Abstract

A key requirement of democratic governance is that policy outcomes and the majority preference of the electorate are congruent. Many studies argue that the more direct democratic a system is, the more often voters get what they want, but the empirical evidence is mixed. This analysis explores the democratic effect of initiatives and referendums theoretically and empirically. The prediction of the formal model is that “bad” representation (i.e., a large preference deviation between the electorate and the political elite) is good for the democratic effect of direct democracy. An empirical investigation of original voter and elite survey data, analyzed with multilevel modeling and post-stratification, supports this argument. Building on the literature, the findings of the analysis suggest that the extent to which direct democratic institutions are conducive for policy congruence—and may thus be advisable as democratic correctives to representative systems—depends on the political conflict structure.

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*“I assume that a key characteristic of a democracy is the continuing responsiveness of the government to the preferences of its citizens” (Dahl, 1971, 1).*

## 1 Introduction

A core element of a democracy is the alignment of legislative and government actions with voters’ preferences. The normative idea that public policies should reflect the majority will of the electorate, what we define as policy congruence, is as a general democratic principle hardly contested. Two literatures approach the study of policy congruence from different angles. First, representation scholars argue that the overlap between the preferences of the electorate and those of the political elite is the key driver of policy congruence (Miller and Stokes, 1963; Golder and Stramski, 2010). Second, the literature on direct democracy argues that referendums and initiatives are conducive for policy congruence because voters are supposed to get what they want when they can participate in policy making (Gerber, 1996; Matsusaka, 2010). However, the empirical evidence on the positive effect of direct democratic institutions on policy congruence is mixed (Lax and Phillips, 2012). For understanding the conditions under which direct democratic institutions are conducive for policy congruence, this analysis focuses on the policy preferences of the electorate and the political elite.

To that end, we provide a formal and empirical analysis. The formal model predicts that the more the preferences of the political elite and the electorate deviate from one another, the larger is the positive effect of initiatives and referendums on policy congruence. Thus, in contrast to the argument of the representation literature, “bad” representation is conducive for the direct democratic effect on policy congruence. Why is that? In essence, initiatives and referendums are powerful democratic correctives when the political elite has clear signals that the voters hold deviating preferences. For the empirical investigation of that argument, we analyze the variation in direct democratic institutions among Swiss cantons (which are in many other dimensions comparable polities). To derive preference measures, we conducted two surveys asking cantonal politicians and voters whether they support a total of 10 tax, health-care, education, and family policies. With this data and fine-grained census information, we estimate elite and voter preference measures, as well as the deviations between the two, using multilevel modeling and post-stratification (Gelman and Little, 1997; Park et al., 2004).

This study makes a substantial contribution to the literature on direct democracy by providing an empirical investigation that relies on consistent measures of policy congruence and accurate estimates of the preferences of the electorate and the political elite over various policy areas. The findings of the empirical analysis support the theoretical prediction that, in direct democratic systems, a large deviation between the preferences of the political elite and the voters is good for policy congruence, not bad, as the representation literature suggests. In a nutshell, direct democratic institutions have no effect on policy congruence when the electorate's and the elites' preferences are aligned; when they deviate, however, referendums and initiatives exert a positive effect on policy congruence, and the conducive effect grows as the elite-voter preference deviation becomes larger. The core substantive contribution of the analysis is that we show how the conducive effect of direct democracy on policy congruence depends on the elite-voter preference deviation, which may, at least partly, explain the mixed empirical findings of the literature.

Last but not least, we discuss the institutional implications of our findings by exploring the political conflict structures under which direct democratic institutions can fully exploit their potential for increasing policy congruence. Our analysis builds on the insights of the representation and the direct democracy literatures, which suggest that competitive and fair elections are powerful in minimizing the elite-voter preference gap on the main political conflict dimension, while direct democratic institutions allow the unbundling of issues and are thus effective democratic correctives for policy questions on second-ordered dimensions ([Besley and Coate, 2008](#)). Accordingly, in a polity with a one-dimensional policy space (e.g., the United States), the conducive effect of direct democracy should be limited because it only matters for the rare policy questions that are not on the main conflict dimension. However, in systems with multiple conflict dimensions, where, for example, the political conflict on cultural policy questions is clearly distinct from the main left-right conflict dimension (e.g., many European countries), we expect strong positive effects, because initiatives and referendums can democratically accommodate a larger number of policy questions on second-ordered dimensions. In sum, the findings of the literature and our analysis suggest that the extent to which direct democratic institutions are advisable as democratic correctives to representative systems depends on the political conflict structure.

The remainder of this article is structured as follows: Section 2 derives the theoretical model; Section 3 discusses the data and the empirical design; Section 4 presents the empirical analysis including instrumental variable regressions; Section 5 discusses the institutional implications of the findings; and Section 6 concludes.

## 2 Theory

We approach the analysis of policy congruence with a theoretical model that shows how the mechanism of direct democracy interacts with the elite-voter preference deviation. The stylized model of policy production analyzes the effects of direct democracy under the scenarios of “good” representation (i.e., a small preference deviation between the elite and the voters) and “bad” representation (i.e., a large deviation). There are a number of formal analyses of direct democracy that explore the initiative and referendum processes in a continuous policy space, whereas the policy can be set arbitrarily far away from an ideal point (Romer and Rosenthal, 1978; Lupia, 1992; Streunenberg, 1992; Gerber, 1996; Hug, 2004). Following Besley and Coate (2008), we adopt a discrete policy space (i.e., a law is either in place or not). The discrete policy space has several advantages: first, the actors’ decision set is reduced to a discrete choice, which is the kind of decision voters face at the ballot box; second, the model can account for referendum and initiative processes; and, third, the discrete policy space allows us to estimate the preferences of the political elite and the electorate on a common scale and thus to derive consistent measures for policy congruence and preference deviations (see Section 4).

The model includes two players, the government ( $G$ ) and the voters ( $V$ ), which both have preferences measured on a continuous scale that represents the degree of support for a specific binary policy decision (e.g., the introduction of a smoking ban).<sup>1</sup> The game starts with a specific law being present or not. The government moves first by choosing between two options: either to maintain the status quo or to propose a policy change. In a second step, the voters either accept the government’s move or, in case they disagree, they may use direct democratic rights (i.e., calling for a referendum, if the government introduced a new law, or launching an initiative,

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<sup>1</sup>We analyze the government as the relevant player of the political elite because we empirically investigate Swiss cantons, where the governments are considered more powerful than the parliaments (Vatter, 2002).

if the government maintained the status quo).<sup>2</sup> If the voters use direct democratic action, the government will lose the vote, and the majority of the electorate will be satisfied. [Figure 3](#) in [Appendix A1](#) shows the game tree of this strategic interaction.

The use of direct democratic rights is associated with costs for the voters, as they have to collect the necessary amount of signatures ( $c_V$ ). In addition, both the government ( $c_G$ ) and the voters ( $c_V$ ) have to pay campaign costs in case of a popular vote.<sup>3</sup> Accordingly, we write the following linear and symmetric utility functions for the government and the voters:

$$\begin{aligned} U_G &= -|x_{out} - x_G| - I_{dd} \cdot c_G \\ U_V &= -|x_{out} - x_V| - I_{dd} \cdot c_V, \end{aligned}$$

where  $I_{dd}$  denotes the use of direct democratic institution ( $I_{dd} = 1$  if the voters call for a referendum or launch an initiative;  $I_{dd} = 0$  if the voters do not take action). The final policy outcome ( $x_{out}$ ) can either be 1 or 0, depending on whether the law is on the books or not. The continuous preferences of the government ( $x_G$ ) and the voters ( $x_V$ ) range from 0 (no support) to 1 (full support). The electorate knows the preference of the government, which acts first and communicates its position in the legislative process. The government, however, does not know precisely the preference of the voters when it decides to change a law, because government officials do not survey constantly the preferences of the electorate, and they cannot fully anticipate the dynamics of a potential campaign ([Romer and Rosenthal, 1979](#); [Matsusaka and McCarthy, 2001](#); [Hug, 2004](#)). Accordingly, the model assumes that the government acts based on a belief of how strong the support in the electorate will be. Technically speaking, the government receives a noisy signal,  $\hat{x}_V$ , which it uses to form a belief about the voters' support for a law. The belief follows a beta distribution (ensuring non-zero probability values between 0 and 1) with an unbiased estimate of the true preference as expected value ( $E(\hat{x}_V) = x_V$ ).<sup>4</sup>

For identifying the conditions of an equilibrium, we assume that a policy does not exist, and we start with the electorate's decision to use direct democratic rights (the last step of the

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<sup>2</sup>Our model only has two players (there is no opposition using direct democratic rights). This simplification would only be problematic if there were no organized political group that would collect signatures for a referendum or an initiative in case the majority of voters disagrees with the government.

<sup>3</sup>The model does not consider electoral costs for the government in taking no action or proposing new laws.

<sup>4</sup>The first shape parameter is a function of  $x_V$  and the second is constant (e.g.,  $s_2 = 30$ ). Accordingly, as  $x_V$  increases, the government assigns higher probabilities to higher values for  $x_V$  ( $\frac{\partial F(x_V)}{\partial x_V} > 0$ ).

game).<sup>5</sup> In case the government prefers the status quo, the voters decide whether they want to launch an initiative for the introduction of a new law.<sup>6</sup> The voters' utility from launching an initiative is  $= -|1 - x_V| - c_V$ , and their utility from not launching an initiative is  $= -|0 - x_V|$ . They will collect signatures for a public vote whenever  $x_V \geq \frac{1+c_V}{2}$  (condition 1); that is, when the support within the electorate ( $x_V$ ) is greater than the costs of launching an initiative ( $c_V$ ). In case the government proposes a new law, the voters decide whether they want to block the legislation with a referendum. The voters' utility from calling a referendum is  $= -|0 - x_V| - c_V$ , and their utility from not calling a referendum is  $= -|1 - x_P|$ . They will collect signatures for a public vote whenever  $x_V \leq \frac{1-c_V}{2}$  (condition 2); that is, when the costs for calling a referendum ( $c_V$ ) are low enough compared to the preference ( $x_V$ ). The bottom line of conditions 1 and 2 is that the electorate weighs its preference against the costs of using direct democratic rights.

The first move in the game—namely, the government's decision to propose a policy change—is a function of the two discussed conditions that specify whether the voters will use direct democratic rights. As discussed earlier, the government acts based on its beliefs about the voters' preference ( $\hat{x}_V$ ). The government's decision of whether it should propose a new policy depends on its own preference for a new law ( $x_G$ ), the costs of a potential public vote ( $c_G$ ), and the anticipated probability that the voters will use direct democratic institutions.<sup>7</sup> A detailed derivation of the inequalities is presented in [Appendix A1](#). Formally expressed, the government introduces a new law, if condition 3 holds:

$$x_G \geq \frac{p_{\text{ref}} \cdot (c_G - 1) - p_{\text{ini}} \cdot (c_G + 1) + 1}{2(1 - p_{\text{ref}} - p_{\text{ini}})},$$

where  $p_{\text{ini}}$  and  $p_{\text{ref}}$  denote the probability that the voters will launch an initiative or call for a referendum, which depends on voters' preference and their costs of using direct democratic rights (conditions 1 and 2). Condition 3 states that the government's decision to introduce a new law is a function of its own and the voters' preferences as well as the usage costs of

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<sup>5</sup>For the equilibrium solutions of the opposite scenario, see [Appendix A1](#).

<sup>6</sup>As in other models of direct democracy, the government always loses in equilibrium when the voters call for a public vote because direct democratic rights are only used in case of an electoral majority (e.g., [Hug, 2004](#)).

<sup>7</sup>The government's belief of  $x_P$  feeds into the expected probability of a referendum. The belief follows a beta distribution and condition 2 ( $x_P \leq \frac{1-c_P}{2}$ ) provides  $p_{\text{ref}} = \int_{x_P=0}^{\frac{1-c_P}{2}} \frac{x_P^{\alpha-1}(1-x_P)^{\beta-1}}{B(\alpha,\beta)}$ , whereas  $\alpha$  and  $\beta$  are the shape parameters and  $B(\cdot)$  is the beta-function.

direct democratic institutions. This condition suggests, for example, that the government may propose laws that deviate from the majority view of the electorate, when the costs for the voters to launch a referendum are high enough.<sup>8</sup> Based on the derived conditions, we investigate how the costs of using direct democratic rights affect the probability that policies are congruent with the majority preference of the median voter, and whether the effect of direct democracy on policy congruence is mediated by the government-voter preference deviation. To that end, we run simulations with the following simulation parameters:

- The government has a strong preference for introducing a new policy ( $x_G = 0.7$ ).
- The support within the electorate for the policy varies between 0.1 and 1.0 ( $x_P \in (0.1, 1.0)$ ).
- The costs of using direct democratic rights are either moderate or high ( $c_V = \{0.35, 0.5\}$ ).
- To account for the government’s probabilistic beliefs regarding voters’ preference, we simulate 1,000 times  $x_P$ . The shape parameters are:  $s_1 = \frac{x_P \cdot s_2}{1 - x_P}$  and  $s_2 = 30$  (see Footnote 4).

Figure 1 plots the comparative statics. The top plot shows how the probability of policy congruence varies as a function of voters’ preferences. The preferences of the median voter and the government are congruent in the area where the majority of the electorate supports the new policy ( $x_V > 0.5$ ). In that area, the law will be introduced and the policy outcome is congruent with the majority preference of the voters (see the 100% probability of policy congruence on the right side of the upper plot). More interesting is the situation in which the preferences of the voters and the government deviate; that is, when  $x_V < 0.5$ . When only a few voters support a new policy, the government will not introduce the policy because it anticipates that the electorate will veto the proposal with a referendum due to the strong popular resistance, and the resulting non-adoption of the policy is congruent with the will of the median voter (see the 100% probability of policy congruence on the left side of the upper plot). However, as the resistance within the electorate decreases (i.e., the voters’ support for the policy increases), the lines start falling: the probability of policy congruence decreases because the voters start refraining from blocking the new law, although the majority still opposes the new law.

For highlighting the effects of direct democracy on policy congruence, we manipulate the costs of using direct democratic rights ( $c_V$ ). The dashed black line shows the findings for a

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<sup>8</sup>Appendix A2 presents the comparative statics and two plots for illustrating the key aspects of the model.

medium level of direct democracy and the gray line for a high level of direct democracy. The dashed line starts falling when the support in the electorate is at about only 20%, while the turning point of the gray line (where  $x_V = \frac{1-c_V}{2}$ ) is at a electoral support of about 30%. Thus, the more direct democratic a system is, the better the voters are shielded from being overruled by a deviating preference of the government.<sup>9</sup> The worst situation, from a policy congruence perspective, is when a bare majority of the electorate are against the government’s proposal. In that case, the voters will not overturn the government’s decision, even when the costs for calling a referendum are low. The lower plot highlights the most interesting range between the turning point of the gray line and the area before the electorate and the government share the same view. The variation of voters’ resistance against the new policy is equivalent to the variation of the government-voter preference deviation (the government’s support for the policy is constant at  $x_G = 0.7$ ). The lower plot shows that as the preference deviation increases, the positive effect of direct democracy on policy congruence becomes larger.

Based on the presented findings, we formulate the following theoretical predictions:

- Policies will be congruent with voters’ majority preference, if the government and the voters share the same preference.
- Direct democratic institutions have a conducive effect on policy congruence, if the preferences of the government and the voters deviate.
- If the preferences of the government and the voters deviate, the positive effect of direct democracy on policy congruence increases as the deviation grows.

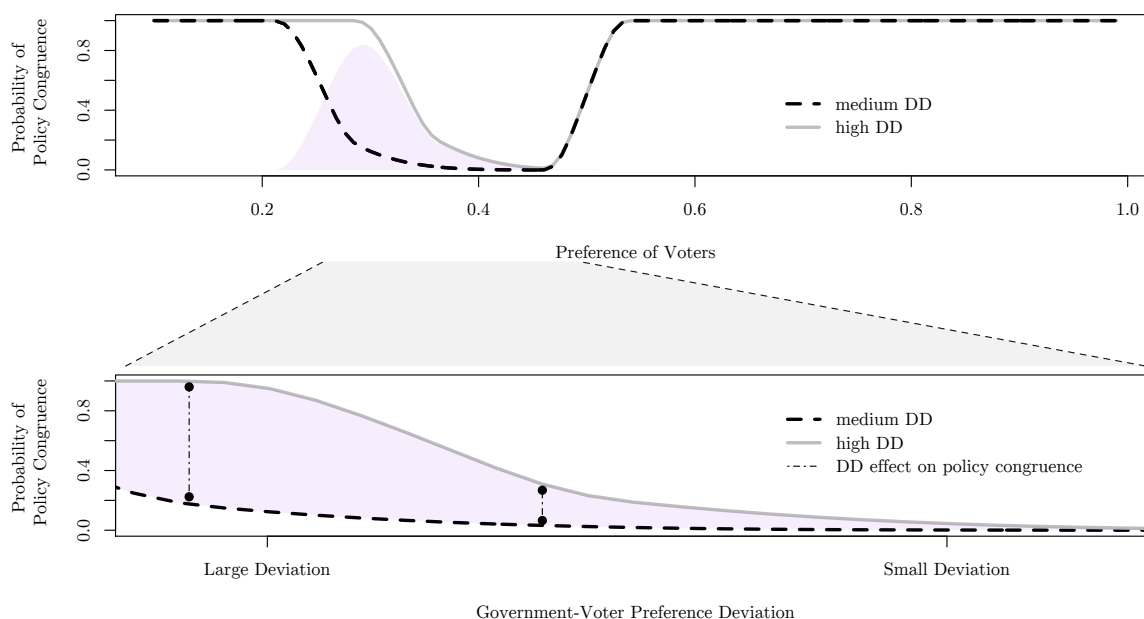
The first two predictions—namely, that policies will be congruent when the elite and the voters share the same preferences and that direct democracy has a positive effect on policy congruence—are quite well established in the literature.<sup>10</sup> Our third prediction, however, that the size of the positive direct democratic effect on policy congruence depends on the elite-voter preference deviation is novel and interesting, as it suggests that in direct democratic

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<sup>9</sup>If the government’s uncertainty about voters’ preferences decreases, the turning points of the lines will remain the same. What changes is that the curves fall and rise faster. In the presented simulation we modeled the second shape parameter to be fairly large ( $s_2 = 30$ ). The plotted curve for medium direct democracy shown in [Figure 1](#) starts falling from 100% predicted policy congruence at the voters’ support of 20% and hits the 0%-probability of policy congruence, where the voters’ support is at 40%. If we decrease the government’s uncertainty, by setting  $s_2 = 1$ , the falling part of the curve is between the more narrow range of 23% to 28% voters’ support.

<sup>10</sup>For an exception, see [Besley and Coate \(2008, 392\)](#), who predict a negative effect of direct democracy on policy congruence under specific circumstances.





Notes: Government preference:  $x_G = 0.7$ ; medium cost (medium DD):  $c_V = 0.5$ ; low cost (high DD):  $c_V = 0.35$ .

Figure 1: Effects of Direct Democracy (DD) and Preferences on Policy Congruence.

systems a large preference deviation is good (not bad) for policy congruence. In other words, the direct democratic mechanism of policy congruence works better under the scenario of “bad” representation (i.e., a large preference deviation), essentially because a large deviation provides clear signals to the government on the preferences of the voters. The mechanics of the theoretical model suggest that the main effect of direct democracy on policy congruence is that the political elite anticipate preference deviations and, in the case that voters have extensive direct democratic rights, follow the position held by (the large) majority of voters. This institutional effect of direct democracy is often referred to as the indirect effect of direct democracy—as opposed to the direct effect, where the actual use of direct democracy drives the policy outcome. A strength of the presented theory is that the key variables (the actors’ preferences and the costs of direct democracy) are operationalizable, and we can thus empirically test the theoretical predictions.

### 3 Measuring Voters' and Elites' Policy Preferences

The key for the empirical analysis of our arguments is that the policies and preferences are measured on the same metric. This is the case in our investigation, as we rely on elite and voter surveys with binary questions that were designed and executed for this research. We asked Swiss voters and politicians whether they support or reject 10 policies that cantons either have or do not have (for the importance of using a binary metric, see [Achen \(1978\)](#); [Erikson et al. \(1993\)](#); [Matsusaka \(2001\)](#); [Hug \(2011\)](#)). The 10 questions cover the most controversial policy questions of the main policy areas in the competence of the cantons, such as tax, family, immigration, education, and health-care policies. The surveys included questions on, for example, the progressivity of cantonal taxes, the voting rights of foreigners, and smoking bans ([Appendix A3](#) lists all policy questions).<sup>11</sup> The broad policy spectrum is important, as it allows us to rule out that our empirical findings are biased because we selected specific policy areas.

We estimate voters' policy preferences on the cantonal level with national survey data from 1,507 respondents. Measures for subnational units cannot be estimated by disaggregating the national survey data, as there are only very few observations for some cantons ([Levendusky et al., 2008](#); [Warshaw and Rodden, 2012](#)). Thus, we take advantage of recent developments in survey research by using multilevel regression and post-stratification (MRP), which provides good estimates, even when the samples for individual subnational units are small ([Gelman and Little, 1997](#); [Park et al., 2004](#); [Lax and Phillips, 2009a,b, 2012](#); [Warshaw and Rodden, 2012](#); [Kastellec et al., 2010](#); [Pacheco, 2012](#); [Tausanovitch and Warshaw, 2013](#)). MRP estimates the preferences of small constituencies in four steps. The first step is to conduct a survey that collects, besides the policy questions, minimal personal information of the respondents (survey step); in the second step, a hierarchical model is fitted to the data to make predictions for specific hypothetical voters (response model step); in the third step, the model estimates are used to make predictions for predefined hypothetical voters (prediction step); finally, based on fine-grained census data, researchers calculate constituency support by weighting the estimated preference of each hypothetical voter according to the number of voters who have the same characteristics as the hypothetical voter in a specific constituency (post-stratification step).

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<sup>11</sup>The survey was conducted in June 2012 with CATI.

For identifying voter categories, we include *gender*, *education* (6 categories), *domicile type* (4 categories), and *age* (4 groups). Altogether, this yields for each canton 192 distinct types of voters ( $2 \times 6 \times 4 \times 4$ ). In the hierarchical model (response model step), we include random effects for the four individual-level variables. Moreover, we add cantonal-level variables, such as the shares of German speakers and of left party support, denoted by the matrix  $\mathbf{X}$ , and we account for regional variation.<sup>12</sup> Finally, we include a cantonal random effect, a fixed effect based on  $\beta\mathbf{X}$ , and a random effect for the region.<sup>13</sup> The model is specified as follows:

$$\begin{aligned}
Pr(y_i = 1) &= \Phi\left(\beta_0 + \alpha_{j[i]}^{gender} + \alpha_{k[i]}^{education} + \alpha_{l[i]}^{domicile} + \alpha_{m[i]}^{age} + \alpha_{n[i]}^{canton}\right) \\
\alpha_j^{gender} &\sim N(0, \sigma_{gender}^2), \text{ for } j = 1, 2 \\
\alpha_k^{education} &\sim N(0, \sigma_{education}^2), \text{ for } k = 1, \dots, 6 \\
\alpha_l^{domicile} &\sim N(0, \sigma_{domicile}^2), \text{ for } l = 1, \dots, 4 \\
\alpha_m^{age} &\sim N(0, \sigma_{age}^2), \text{ for } m = 1, \dots, 4 \\
\alpha_n^{canton} &\sim N(\alpha_{o[n]}^{region} + \beta\mathbf{X}_n, \sigma_{canton}^2), \text{ for } n = 1, \dots, 26 \\
\alpha_o^{region} &\sim N(0, \sigma_{region}^2), \text{ for } o = 1, \dots, 7
\end{aligned}$$

MRP takes advantage of the data structure, as we estimate a hypothetical voter’s predicted support for a policy, based on all the people in the sample who share her attributes. Concretely, when we estimate whether a 55-year-old woman with a university degree, who lives in a rural village, supports a policy or not, we use information from the entire sample, but the estimates are especially influenced by the answers from the 50- to 64-year-old people, the university degree holders, the people living in rural areas, and the women. The property of partial pooling is key for the accuracy of MRP, as we rely neither solely on the average for all voters of a canton, nor on the average of the entire sample, but rather on the weighted mean of the two averages (Steenbergen and Jones, 2002; Gelman and Hill, 2007).

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<sup>12</sup>*Education* categories: 1 (mandatory schooling or no response), 2 (apprenticeship), 3 (university-entrance diploma [Matura], teachers college), 4 (additional job training [höhere Fachausbildung]), 5 (Advanced training [Höhere Fachhochschule]), 6 (university degree including U. of App. Sciences); *Domicile* categories: 1 (urban center), 2 (agglomeration), 3 (isolated city), 4 (rural area); *Age* categories: 1 (18-34 years), 2 (35-49 years), 3 (50-74 years), 4 (75- years); *Region* categories: 1 (Geneva, Valais, Vaud), 2 (Bern, Fribourg, Jura, Neuchâtel, Solothurn), 3 (Aargau, Basel-Stadt, Basel-Landschaft), 4 (Zurich), 5 (Appenzell I. Rh., Appenzell A. Rh., Glarus, Grisons, St. Gallen, Schaffhausen, Thurgau), 6 (Lucerne, Nidwalden, Obwalden, Schwyz, Uri, Zug), and 7 (Ticino).

<sup>13</sup>To specify the best (not over-fitted) response model, we estimated for each policy question 64 combinations of five different cantonal predictors and chose the models that minimize AIC and BIC. For more detailed information on the model specifications, see Appendix A4.2.

Using the model estimates, we carry out steps 3 and 4 (prediction for all hypothetical voters and post-stratification with census data). We have 192 different hypothetical voters in each canton. For each of these hypothetical voters, the model predicts the probability with which these voters support a policy. Since we have in addition information about the variation among cantons, we create predictions ( $\hat{\pi}_{ng}$ ) for all hypothetical voters in each canton. In the final step (post-stratification), we sum the predicted support among all hypothetical voters and weigh each type by its frequency in a given canton ( $N_{ng}$  with data from the census ( $N_n$  is the total number of eligible voters):

$$\hat{\pi}_n = \frac{\sum_{g \in n}^G \hat{\pi}_{ng} N_{ng}}{N_n}$$

Following these four steps, we estimate the cantonal support of the electorate for each policy. Besides the cantonal preference of voters, we also estimate the preference of the elite (i.e., the cantonal governments) with the same method. To that end, we conducted an online survey with cantonal politicians. The survey asked cantonal politicians about their preferences for the 10 policies under investigation. We selected politicians from cantonal parties that represent more than 10% of the electorate. The final sample consists of 431 respondents. Thanks to a targeted follow-up survey, the sample includes at least three respondents for each cantonal party in the sample.<sup>14</sup> We again apply the four steps of the MRP method to derive cantonal elite preferences using the elite survey data on the policy preferences of the cantonal politicians. However, for the estimation of the elite preferences, we use a hierarchical model specification and weights for the post-stratification that are tailored to the application of MRP for cantonal elite preferences. The elite hierarchical model includes the party affiliation of cantonal politicians as an individual-level variable, which is a very strong predictor of whether a cantonal politician supports a policy or not. In addition, we again model cantonal-level variables and random effects for the cantons

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<sup>14</sup>For the first online survey wave on March 15, 2013, we invited 1,046 cantonal politicians with personalized e-mails to participate in the survey. After analyzing the sample of the first wave, we contacted, on March 28, 2013, 99 additional cantonal politicians from the yet underrepresented cantonal parties (again with personalized e-mails). Of the total 476 received responses, we dropped 45 from the sample as they either included impossible canton and party combinations or were double entries. The final sample consists of 431 respondents (38% of the 1,145 contacted politicians).

and for the regions in the response model.<sup>15</sup>

In the next (third) step, we predict for each cantonal party the policy preferences with the specification of the elite hierarchical model. Finally, in the post-stratification step, we derive aggregated cantonal elite preference measures that are representative of the political strength of the relevant parties in each canton. To that end, we weigh the predicted cantonal party preferences with the government compositions of each canton. In short, we estimate cantonal elite preferences for all 10 policies by first predicting policy support within each relevant cantonal party based on our elite survey data, before aggregating the estimated party preferences to a cantonal elite preference measure using government composition data. [Appendix A4](#) provides more detailed information on how we derive the preference measures for the voters and the elite with MRP.

Our empirical design overcomes three major problems of the literature on the institutional effects of direct democracy. First, most articles investigate whether specific policies (e.g., tax levels) correlate with variation in direct democracy ([Matsusaka, 2004](#); [Feld and Kirchgässner, 2000](#)). The validity of these analyses hinges on the implicit (and implausible) assumption that voters' preferences are constant between the units of analysis. Second, studies analyzing voters' preferences typically measure the preferences and policy outcomes on different scales ([Achen, 1978](#); [Erikson et al., 1993](#); [Matsusaka, 2001](#); [Hug, 2011](#)). Third, many articles investigate only a single or a few policy areas, which might bias the findings ([Gerber, 1996](#); [Lax and Phillips, 2009a](#)). We overcome the mentioned problems by estimating, for various policy areas, the preferences of the voters and the elite on the same (binary) metric as the policy outcomes. The measures for policy congruence and for the elite-voter preference deviation are consistent because of the common binary metric.

## 4 Empirical Analysis

The core idea of policy congruence is that the actions of representatives are in line with the majority positions of voters. Following our definition of policy congruence, we analyze an out-

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<sup>15</sup>We chose the combination of explanatory variables on the cantonal level in the response model (**X**) that minimize AIC and BIC (see [Footnote 13](#)).

come variable that codes whether a policy in a given canton is supported by the majority of the citizens or not (Lax and Phillips, 2009a, 2012). We rely on this binary measure of policy congruence for conceptual and methodological reasons. Conceptually, policy congruence is not about the degree of support for a policy in the electorate. Rather, the basic question is whether a majority of voters is in favor of a policy or not.<sup>16</sup> Methodologically, the dichotomy of the policies (i.e., they are either in place or not) is important because this allows us to measure policies and preferences on a common (binary) metric. This common metric is also reflected in the binary coding of policy congruence, which essentially combines the measures of the preferences of voters and the policy outcomes. Of course, the binary measurement of policies and policy congruence simplifies political realities. However, we believe that the costs of reducing complexity are more than compensated by the gains in conceptual clarity and methodological rigor.

The two main explanatory variables of the theoretical model presented in Section 2 are the costs of using direct democratic rights and the elite-voter preference deviation. To measure the costs of direct democracy usage, we rely on a widely used index of direct democratic rights (Frey and Stutzer, 2000; Freitag and Vatter, 2006; Stadelmann-Steffen and Vatter, 2011; Stadelmann-Steffen and Freitag, 2011). The index measures how difficult it is to use direct democracy, for example, in terms of the time granted to collect signatures and the number of signatures required (Stutzer, 1999; Frey and Stutzer, 2000). In line with the theoretical model, the direct democracy measure accounts for the usage costs for citizens to successfully call for a referendum or to launch an initiative (high values on the index imply low costs and thus more direct democracy). Appendix A5 shows the variation of the direct democracy index and compares the values of the Swiss cantons to the level of direct democracy in California. In comparative perspective, Swiss cantons vary between a low level of direct democracy and a very high level. The elite-voter preference deviation is measured as the absolute distance between the government's and the voters' preferences in case they hold opposing majority opinions (if they agree, the variable is set to 0).

As control variables, we use measures of the clarity of the government and voter opinions,

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<sup>16</sup>We coded whether a policy was cantonal law when the surveys were conducted (that is, at the end of 2012).

coded as the absolute deviation from 0.5, because the findings on the preference-deviation variable could be driven by extreme voter or government opinions (Lax and Phillips, 2009a). In addition, we control for a number of institutional, economic, cultural, and structural variables. We include measures of the electoral threshold and the power of the legislator vis-à-vis the government, expecting that low electoral thresholds and powerful legislators are associated with higher policy congruence (Kaiss, 2010; Bochsler and Bousbah, 2015). Furthermore, we control for economic, cultural, political, and structural heterogeneity among the cantons by modeling GDP per capita, population size, the strength of the Swiss People’s Party (SVP), and the share of German speaking citizens. Particularly, the language variable is of importance because direct democracy is more comprehensive in the German speaking part of Switzerland. Finally, we model a dummy variable that codes for each canton and policy whether there has been a public vote or not in the three years prior to the surveys (that is, 2010 – 2012).<sup>17</sup> This direct democracy in use measure is important because scholars differentiate between the effects of direct democracy in use (policies are decided on the ballot) and the effects of the threat of a possible referendum or initiative on the behavior of legislators. As it is conventional in the literature and consistent with our theoretical model, we rely on an institutional measure on the usage costs of direct democratic rights as our main explanatory variable (Matsusaka, 1995; Gerber, 1996; Stutzer, 1999). However, including the dummy variable on the actual use of direct democracy allows us to analyze whether policies are primarily congruent because voters have made decisions at the ballot box.

We analyze cross-nested hierarchical probit models with random effects for policies and cantons. The first model includes the institutional direct democracy measure. The estimates reported in Table 1 confirm the positive and significant effect of direct democratic institutions on policy congruence. In the second model, we introduce the preference deviation variable. The findings show, as suggested by the work of the representation literature, that preference deviation is a significant predictor of policy incongruence: the more the people and the government disagree, the less likely is policy congruence. In the third model, we empirically test the main prediction of the theoretical model—namely, that the positive effect of direct democracy

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<sup>17</sup>Own coding based on data from the Centre for Research on Direct Democracy (c2d). The findings are the same if we extend the lag to five or more years.

increases, as the preference deviation between the voters and the government becomes larger. To investigate that prediction, we include an interaction between the preference deviation and direct democracy. Model 3 includes all three explanatory variables that are motivated by the theoretical model. As expected, the interaction term is positive and significant. This finding supports our main prediction that the positive effect of direct democracy increases, the more the preferences of the voters and the elite are at odds (a hypothesis test for the significance of the interaction term is presented below). The empirical findings also confirm the prediction that direct democratic institutions have no effect on policy congruence, if the government and the people both support or reject a policy (i.e.,  $\text{DEVIATION} = 0$ ).

In order to test our main hypothesis against alternative explanatory factors, the full Model 4 includes all controls as well as the direct democracy in use variable. The results of Model 4 are essentially identical to the estimates of Model 3. Most importantly, the positive and significant effect of the interaction term corroborates the main prediction of the model that the conducive effect of direct democracy on policy congruence increases as a function of the degree of the dissonance between the government and voter preferences. Model 4 further shows that government and voter opinion clarities are significant predictors of policy congruence, but controlling for extreme voter or government opinions does not alter the findings on the preference-deviation variable. The only other significant predictor is the share of German speakers (policy congruence is more widespread in German speaking cantons). Finally, the estimates of the direct democracy in use variable suggest that it does not matter for policy congruence whether the electorate has voted on an issue or not (note that the model also includes the usage costs of direct democratic rights and the interaction effect). According to this non-finding, policy congruence is not primarily driven by the usage of direct democracy, which is consistent with both our theoretical model and the conventional view in the literature (Matsusaka, 1995; Gerber, 1996). Appendix A6 reports robustness checks using alternative measures of direct democracy and of the preference deviation between the government and the voters. The results for the main explanatory variables are robust.

Besides the robustness to alternative specifications, a further concern related to the validity of the presented findings is the potential endogeneity that policy congruence influences (in



	Model 1	Model 2	Model 3	Model 4
DIRECT DEMOCRACY	0.19** (0.09)		0.12 (0.10)	-0.20 (0.17)
GOVERNMENT-VOTER PREFERENCE DEVIATION		-2.57** (1.00)	-14.17*** (5.32)	-14.71*** (5.52)
DD x PREFERENCE DEVIATION			2.46** (1.10)	2.72** (1.15)
VOTER OPINION CLARITY				3.44** (1.52)
GOVERNMENT OPINION CLARITY				-2.15** (1.08)
GDP (PER CAPITA IN 100K)				0.00 (0.54)
ELECTORAL THRESHOLD				-0.01 (0.01)
% GERMAN SPEAKERS				1.16** (0.55)
LEGISLATIVE POWER				-0.06 (1.05)
POPULATION SIZE (IN 100K)				-0.02 (0.04)
SVP VOTE SHARE				-0.01 (0.02)
DD IN USE (PAST 3 YEARS)				-0.29 (0.27)
CONSTANT	-0.81* (0.41)	0.14 (0.19)	-0.36 (0.45)	0.52 (1.19)
<i>CPC</i>	71%	68%	69%	72%
BIC	360.19	359.21	358.65	390.44
<i>ll</i>	-168.97	-168.48	-162.64	-153.52
<i>N</i>	260	260	260	260
Groups: Cantons	26	26	26	26
Groups: Policies	10	10	10	10
Variance: Canton	0.03	0.07	0.04	0.00
Variance: Policy	0.33	0.25	0.26	0.19

*CPC*: Correctly predicted cases; baseline is 51%. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

Table 1: Cross-Nested Hierarchical Probit Model Estimates of the Effects of Direct Democracy and Government-Voter Preference Deviation on Policy Congruence.

reverse model direction) the extent of direct democracy in the cantons, which would invalidate the above discussed inferences. To out rule potential endogeneity, we run 2SLS IV estimation using the degree of direct democracy at the end of the 19th century as an instrument for direct democratic institutions in 2010. The analysis relies on the Historical Direct Democracy Index, which dates back to 1803 (Leemann, 2014). The instrument fulfills the exclusion restriction, when the degree of direct democracy in the 19th century is independent from the error term of the second-stage regression. This assumption is violated in case of reverse causation, or if the

historical institutions affect the outcome through another causal pathway than the instrumented variable. The former is unproblematic because policy congruence today cannot be the cause of institutional settings of the 19th century. The latter is relevant when the historical institutions affect a variable that is not modeled in the second stage, but influences the outcome even after controlling for all other explanatory factors. We cannot think of such an alternative causal pathway.

To assess the strength of the instrument, we use the  $F$ -statistic of the instrument in the first stage model. The direct democracy index from the end of the 19th century achieves an  $F$  value of 12.45, which is clearly higher than the conventional threshold of 10 (Stock et al., 2002).<sup>18</sup> We perform 2SLS IV estimation for the theoretically motivated Model 3 as well as for the full Model 4 of Table 1 (Angrist and Pischke, 2008). Table 2 reports the findings of the first- and second-stage models. The estimates are by and large the same as the ones discussed above. Most importantly, the interaction effect between direct democracy and the government-voter preference deviation is clearly positive and significant in both models, corroborating the main argument of this analysis.

Figure 2 illustrates the size of the interaction effect relying on the estimates of Model 3 of Table 1. Reported are the predicted probabilities of policy congruence, including the uncertainty of the predictions (Herron, 1999; Gelman and Hill, 2007). The  $x$ -axis displays the government-voter preference deviation and the  $y$ -axis the predicted probability of policy congruence. The solid line reports the predictions for a very low level of direct democracy (limited DD) and the dotted line for a hypothetical canton with high values of direct democracy (extensive DD). The increase in direct democracy from limited to extensive has two effects: first, citizens in cantons with more direct democracy more often get what they want. The probability of policy congruence is always higher, no matter how large the preference deviation is. For example, for a small value of the preference deviation (0.00), the difference in predicted probabilities of policy congruence between limited and extensive direct democracy is about 16.8% (95% CI: [-0.03, 35.18]), while this measure increases to 40.7% (95% CI: [16.86, 58.28]) for a slightly more

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<sup>18</sup>We account for the uncertainty of the first-stage model in the estimation by simulating 500 first-stage predictions, which are then translated into second-stage estimations. Because the canton of Jura did not exist before 1973, we cannot use the instrument for this case and have therefore dropped the canton from the analysis.

	Model 6 1st Stage	Model 3 IV 2nd Stage	Model 4 IV 2nd Stage
DIRECT DEMOCRACY (19TH CENTURY)	0.43*** (0.12)		
GOVERNMENT-VOTER PREFERENCE DEVIATION	2.16** (0.98)	-15.67** (6.10)	-15.00** (6.13)
DIRECT DEMOCRACY		0.22* (0.13)	0.02 (0.19)
DD x PREFERENCE DEVIATION		2.79** (1.27)	2.79** (1.28)
VOTER OPINION CLARITY			3.19** (1.56)
GOVERNMENT OPINION CLARITY			-1.51 (1.12)
GDP (PER CAPITA IN 100K)			0.24 (0.64)
ELECTORAL THRESHOLD			-0.01 (0.01)
% GERMAN SPEAKERS			0.73 (0.52)
LEGISLATIVE POWER			0.02 (1.07)
POPULATION SIZE (IN 100K)			-0.00 (0.03)
SVP VOTE SHARE			-0.00 (0.02)
DD IN USE (PAST 3 YEARS)			-0.26 (0.27)
CONSTANT	2.54*** (0.41)	-0.78 (0.58)	-0.59 (1.35)
R <sup>2</sup>	0.48		
BIC		340.71	379.36
<i>ll</i>		-153.79	-148.17
<i>N</i>	25	250	250
Groups: Cantons		25	25
Groups: Policies		10	10
Variance: Canton		0.03	0.02
Variance: Policy		0.30	0.28

\*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

Table 2: 2SLS IV Probit Model Estimates with 19th Century Institutions as Instruments.

sizable preference deviation (0.10).

Second, the differences in the predicted probabilities also document the interaction effect between direct democracy and the preference deviation: the difference in the predicted probabilities between limited and extensive direct democracy increases as the deviation becomes larger (that is, the effectiveness of adding direct democratic rights on policy congruence grows, the larger the preference deviation). The interaction effect is clearly significant, as the confi-

dence interval of the difference in the differences shows ( $\Delta(\Delta\hat{p}) = 24.38\%$  (95% CI: [2.52, 43.42]) (Brambor et al., 2006; Tsai and Gill, 2013). In short, the robust and significant results support the main prediction of our theoretical model—namely, that the positive effect of direct democratic institutions increases, as the elite-voter preference deviation becomes larger.<sup>19</sup>

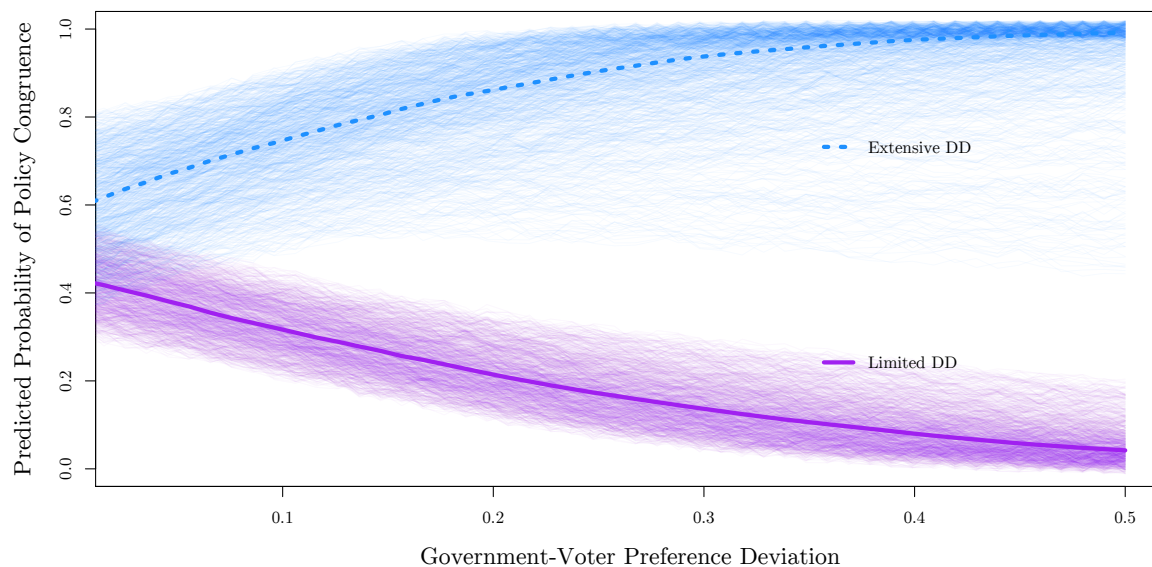


Figure 2: Interaction Effect of Direct Democracy and the Government-Voter Preference Deviation on Policy Congruence.

Figure 2 also shows that the effect of the preference deviation on policy congruence turns from positive to negative when we change the decision making from strongly to limited direct democratic (the line goes down). This may need further explanation. In general, a large preference deviation between the elite and the citizens means that the preference of the electorate is not well represented in policy making. Bad representation is bad for policy congruence. The model-based predictions for a system with very limited direct democratic rights shows the dominance of that effect. The slope falls because the negative effect of the preference gap on policy congruence dominates. In other words, the costs of using direct democracy become too

<sup>19</sup>A potential objection against the hypothesis tests is that we underestimated the full uncertainty in Models 1 to 4 because some explanatory variables are measured with uncertainty (the preference variables) (see, e.g., Lewis and Linzer, 2005). To account for the full uncertainty, we estimate the same models relying on 1,000 posterior draws of the preference predictions (first stage), rerun the models for each draw, and save 30 draws of the second-stage posterior. The resulting sum of draws reflects both the uncertainty from measurement (first stage) and the variation in the estimated coefficients (second stage). This approach still yields a significant difference of differences estimate (on the 0.05  $\alpha$ -level).

high to correct the growing preference gap. This does not apply, however, for a political system with a high level of direct democracy. In such a system, a large elite-voter preference deviation is conducive for policy congruence because the elites know that voters are so strongly opposed to their position that the intervening factors (i.e., the imprecision of the preference signal and the usage costs of direct democratic institutions) become inconsequential. Thus, in strongly direct democratic polities, a large elite-voter preference deviation empowers the electorate to overturn the deviating preferences of the elite.

## 5 Direct Democracy and the Political Conflict Structure

The finding that the effect of direct democracy on policy congruence depends on the preferences of the elite and the voters may explain why some scholars have found evidence that direct democracy leads to more policy congruence (Gerber, 1996, 1999; Matsusaka, 2004, 2010), while others could not replicate that result (Lascher et al., 1996; Lax and Phillips, 2009a, 2012; Tausanovitch and Warshaw, 2014). Direct democracy has no unique and constant effect on policy congruence. The main finding of our analysis—that direct democracy becomes more conducive for policy congruence the more the preferences of the people and the elite deviate—raises the question of when we would expect large preference gaps and thus strong effects of direct democracy.

Representation scholars argue that competitive and fair elections constantly minimize the elite-voter preference gap through selection and sanctioning: citizens (s)elect representatives with similar preferences and sanction incumbents by not re-electing them if their behavior in office deviates from the majority will of the electorate (Schumpeter, 1942; Mansbridge, 2009). Scholars of direct democracy, however, have emphasized that this electoral mechanism cannot democratically accommodate all policies in a multidimensional policy space. Here lies the powerful property of direct democracy: referendums and initiatives allow it to unbundle policy issues (Besley and Coate, 2008). In other words, citizens can also influence the outcomes of policies that are not democratically accommodated by elections, when they live in representative democracies that are complemented with direct democratic institutions.<sup>20</sup>

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<sup>20</sup>Direct democratic systems and representative democracies are often described as distinct models (Held,

The effectiveness of the direct democratic and the representative mechanisms of policy congruence thus depend on the political environments. Politics in modern democracies is typically dominated by a main political conflict dimension on the left-right spectrum. In a polity with a one-dimensional political conflict structure, competitive and fair elections should regularly minimize the elite-voter preference gap for the overwhelming majority of policies. Adding direct democratic institutions in that case would only matter for the rare policies not covered by the main political conflict dimension. Accordingly, the effects of direct democracy on policy congruence should be limited because we expect only rare and small elite-voter preference gaps. An example of a polity with a one-dimensional policy space is the United States in the last 130 years, where a large share of issues aligned on the liberal-conservative dimension (McCarty et al., 2006).

Unlike in the United States, political conflict in most European countries spans over a multidimensional policy space (Kriesi et al., 2008).<sup>21</sup> The dominant (first) dimension usually covers the classic left-right economic conflict, while the second-ordered dimensions include, among others, moral and cultural policy questions (Kitschelt and Rehm, 2014). If parties compete primarily along the economic (distributional) dimension, we would expect elections to accommodate the policies on the first dimension. In that case, direct democratic institutions should be powerful in aligning second-ordered policies with the majority preference of the electorate. Taken together, the insights of the representation literature and the unbundling argument suggest that direct democratic institutions are effective democratic correctives for policy questions not covered by the dominant (first) dimension.

We find empirical evidence supporting the argument that the effectiveness of direct democracy depends on the political conflict structure. First, the majority of initiatives raised over the last 50 years in Switzerland were concerned with policies on the second-ordered dimension

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1996). Kriesi (2005, 6), however, argues that for empirical research, this distinction is “overdrawn” because we usually analyze democratic polities that are enriched with direct democratic elements. This is also true for Swiss cantons, which are representative systems complemented to varying degrees by direct democratic institutions (see Appendix A5). Most scholars argue that adding direct democratic institutions to representative democracies fosters the participatory capacity of voters (Kriesi, 2005; Matsusaka, 2005). Stadelmann-Steffen and Freitag (2011, 527) make the contrasting point that the combination of representation and direct democracy “offers the least advantageous environment for the development of civic engagement.”

<sup>21</sup>We rely on this stylized distinction between one- and two-dimensional polities for clarifying the argument. Political competition occurs on multiple dimensions in all democratic systems.

(Leemann, 2015). Second, according to the data of this study, the preference gap between the elite and voters is larger for policies not covered by the (dominant) first dimension. Third, the analysis of the data also suggests that the effect of direct democracy on policy congruence is driven by the policies that are not covered by the first dimension. As a cautionary note, we would like to emphasize that the dimensionality analysis of our data relies on the strong assumption that the political conflict structure is the same in all the 26 cantons. The findings are thus more suggestive than conclusive (see [Appendix A7](#) for a detailed discussion of the dimensionality analysis and its limitations).

In sum, we conclude that the insights of previous research and the analysis of our data point in the same direction: whether it is—from a policy congruence perspective—advisable to introduce direct democratic institutions as democratic correctives to representative systems depends on the political conflict structure (i.e., the dimensionality of the political space). This broader institutional implication of our analysis illustrates that direct democratic institutions interact in a complex way with the key elements of representative democracies.

## 6 Conclusion

This article presented a theoretical model of how direct democratic institutions achieve policy congruence. The formal model showed that the positive effect of direct democracy on policy congruence increases, as the elite-voter preference deviation grows. This finding suggests, in contrast to the argument of representation scholars, that “bad” representation is in direct democratic systems conducive for policy congruence, mainly because it provides a clear signal that the elite and the voters disagree on some policy issue. For the empirical analysis, we relied on original elite and voter survey data for 10 different policies of Swiss cantons, which provide unique variation on direct democratic institutions. The specification of the empirical model followed directly from the theoretical analysis, and the empirical results were as predicted and robust across different specifications (also in models with an instrument for the institutional variable). To the best of our knowledge, this analysis is the first policy congruence study that investigates theoretically and empirically the connection between direct democratic institutions and the preferences of the voters and the elite over multiple policy areas.

Some points of concern need to be discussed. Objections against the formal model might be that the equilibrium predicts the absence of referendums and that the government always loses direct democratic votes. Both are obviously not the case in the real world. However, we believe that this does not limit the insights the model provides in respect to the direct democratic mechanism of policy congruence we have analyzed. According to the theoretical model and empirical findings, the main democratic power of direct democracy is that the political elite anticipate preference deviations and, in the case that voters have extensive direct democratic rights, follow the position held by (the large) majority of voters. This is often referred to as the indirect effect of direct democracy. Of course, policy outcomes can also become congruent through the use of direct democratic rights (that is, the direct effect of direct democracy). However, the empirical results, in support of the theoretical model, do not suggest that this direct effect of direct democracy on policy congruence is strong.

The model's discrete policy space might be subject to criticism too. We argue that for direct democratic decision making, this simplification is appropriate because voters face a discrete choice at the ballot box. Nevertheless, we would like to emphasize that the political elite has non-binary room to maneuver for shifting away from voters' preferences, particularly in the implementation of laws. The discrete policy space might thus be an inappropriate simplification for some stages of the policy cycle, like the implementation stage, but less so for the decision-making part we analyzed in this study. In any case, the fact that the political elite has some discretion in the implementation of policies highlights the limitations of direct democratic control (Gerber et al., 2000). A further question in that regard is to what extent well-organized and heavily funded interest groups control the use of direct democratic institutions and the content of direct democratic campaigns. The more this is the case, the less accurate is our theoretical distinction between the preferences of the elites and the voters. This is less problematic in the small-sized Swiss cantons, where these kind of interest groups hardly exist, as compared to U.S. states like California (Broder, 2000).

Hug (2009) forcefully argues in a recent critique of the literature on direct democracy that further research is needed to better understand how direct democratic institutions interact with the key elements of representative democracy. As far as policy congruence is concerned, we

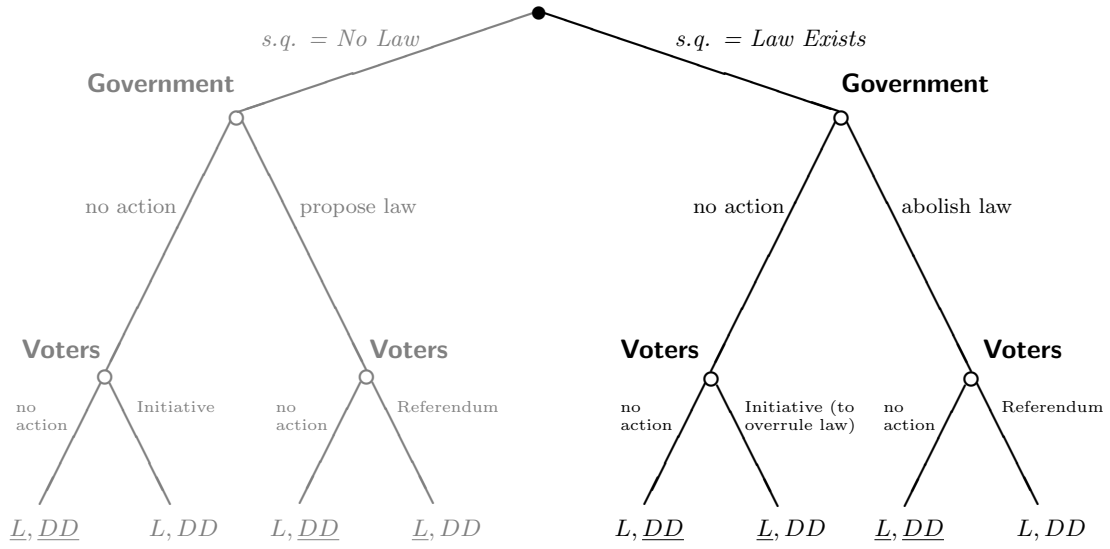


believe our analysis provides a nuanced account of how adding participatory elements to representative democracies affects democratic performance. The presented findings show that the extent to which direct democratic institutions lead to more policy congruence depends on the elite-voter preference deviation. More specifically, the democratic effect of direct democracy is most effective in cases of “bad” representation. Finally, the discussion of the institutional implications of the findings suggests that initiatives and referendums are most likely effective democratic correctives to representative systems in polities with multidimensional conflict structures.

# A1 Formal Analysis

## A1.1 Supplements to the Theoretical Analysis of Section 2

Figure 3 illustrates the setup of the model discussed in Section 2, where we analyze the following status quo (s.q.): there is no law and the government either proposes a new law or takes no action (see gray part of Figure 3). The next subsection shows the mirror equilibrium analysis for the case that the law exists and the government either takes no action or abolishes the law (see black part of Figure 3).



Notes:  $L$  = a law is present,  $\underline{L}$  = there is no law;  $DD$  = an initiative or a referendum is used,  $\underline{DD}$  = there is no use of direct democratic institutions;  $s.q.$  = status quo.

Figure 3: Theoretical Model: Extensive Form

Section 2 discusses condition 3, which formalizes the government's decision to propose a new law, without providing the formal derivation of the inequality condition. Condition 3 is derived from the government's expected utilities as follows:

$$\begin{aligned}
 E[U_G(\text{proposing } L)] &\geq E[U_G(\text{no action})] \\
 p_{\text{ref}} \cdot U_G(\underline{L}|DD) + (1 - p_{\text{ref}}) \cdot U_G(L|\underline{DD}) &\geq p_{\text{ini}} \cdot U_G(L|DD) + (1 - p_{\text{ini}}) \cdot U_G(\underline{L}|\underline{DD}) \\
 p_{\text{ref}} \cdot (-|0 - x_G| - c_G) + (1 - p_{\text{ref}}) \cdot (-|1 - x_G|) &\geq p_{\text{ini}} \cdot (-|1 - x_G| - c_G) + (1 - p_{\text{ini}}) \cdot (-|0 - x_G|) \\
 x_G &\geq \frac{p_{\text{ref}} \cdot (c_G - 1) - p_{\text{ini}} \cdot (c_G + 1) + 1}{2(1 - p_{\text{ref}} - p_{\text{ini}})}
 \end{aligned}$$

## A1.2 Mirror Equilibrium

We explore in the following the mirror equilibrium analysis for the case that the law exists (see black part of [Figure 3](#)), assuming that the government does not change the status quo and that the voters can either accept the law or launch an initiative. To find the conditions for **an equilibrium**, we again start with the last node in the game tree. The voters launch an initiative, if the following condition 4 holds (equivalent to condition 2 in [Section 2](#)):

$$\begin{aligned}
 U_V(\underline{L}|DD) &\geq U_V(L|\underline{DD}) \\
 -|0 - x_V| - c_V &\geq -|1 - x_V| && \text{Note: } 0 \leq x_G, x_V \leq 1 \\
 -x_V - c_V &\geq x_V - 1 \\
 x_V &\leq \frac{1 - c_V}{2}
 \end{aligned}$$

In case the government decides to abolish the law, the voters call for a referendum, if condition 5 holds (equivalent to condition 2 in [Section 2](#)):

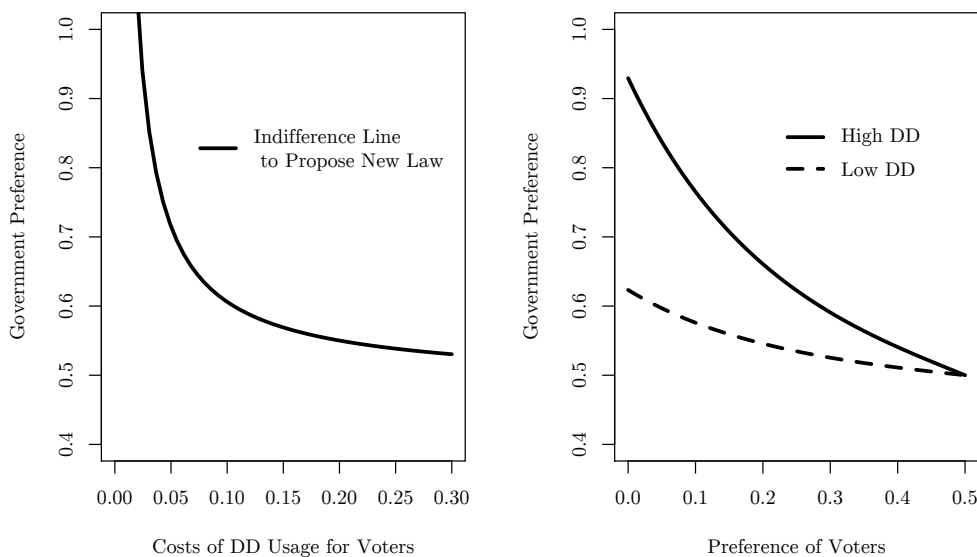
$$\begin{aligned}
 U_V(L|DD) &\geq U_V(\underline{L}|\underline{DD}) \\
 -|1 - x_V| - c_V &\geq -|0 - x_V| && \text{Note: } 0 \leq x_G, x_V \leq 1 \\
 x_V - 1 - c_V &\geq -x_V \\
 x_V &\geq \frac{1 + c_V}{2}
 \end{aligned}$$

The government can either propose to change the law or stick with the status quo. Either way, the government has to take into account the possibility that the electorate will overturn its decision with the use of direct democratic rights. For both actions (keeping or abolishing the law) the government forms expected utilities with some uncertainty about the exact position of the median voter's ideal point. Based on these expected utilities, we derive the following condition 6, which formalizes the government's decision (and is equivalent to condition 3, see [Subsection A1.1](#)):

$$\begin{aligned}
 E[U_G(\text{keeping } L)] &\geq E[U_G(\text{abolishing } L)] \\
 p_{\text{ini}} \cdot U_G(\underline{L}|DD) + (1 - p_{\text{init}}) \cdot U_G(L|\underline{DD}) &\geq p_{\text{ref}} \cdot U_G(L|DD) + (1 - p_{\text{ref}}) \cdot U_G(\underline{L}|\underline{DD}) \\
 p_{\text{ini}} \cdot (-|0 - x_G| - c_G) + (1 - p_{\text{ini}}) \cdot (-|1 - x_G|) &\geq p_{\text{ref}} \cdot (-|1 - x_G| - c_G) + (1 - p_{\text{ref}}) \cdot (-|0 - x_G|) \\
 x_G &\geq \frac{p_{\text{ini}} \cdot (c_G - 1) - p_{\text{ref}} \cdot (c_G + 1) + 1}{2(1 - p_{\text{ref}} - p_{\text{init}})}
 \end{aligned}$$

## A2 Comparative Statics

Figure 4 plots comparative statics based on the theoretical model, illustrating how the costs of direct democratic action affect each player's decisions. Let us start with the left plot. The  $y$ -axis shows the government's support for a new law and the  $x$ -axis voters' costs to launch a campaign. The government's decision to propose a new law is a function of its own support for the reform and the direct democratic usage costs for voters. When the electorate can easily call for a referendum (i.e., the costs of using direct democratic rights for voters are small), the government will only enact new laws that it strongly supports. But as the costs for the voters increase, the government will propose a law that is supported by only a small majority within the government. Another way to illustrate the effect of direct democratic institutions is shown in the right plot. The  $y$ -axis plots again the government support for a new law and the  $x$ -axis the median voter's preference. The stronger the electorate's supports for a bill (although overall still rejecting it, as  $x_P < 0.5$ ), the more inclined the government is to pass the law. The comparison between the solid and the dashed line shows that the government is more responsive to the median voter's preference, if the usage costs for direct democracy are lower (see comparison between solid and dashed lines).



Notes: Parameter values: Left:  $x_P = 0.1$ ,  $c_G = 0.02$ . Right:  $c_P = 0.1/0.3$ ,  $c_G = 0.05$ .

Figure 4: Comparative Statics Illustrating How Direct Democracy Affects Player's Decisions.

## A3 Policy Items and Responses (our translation)

### Tax Policies

Policy 1 The canton grants special tax cuts for wealthy foreigners.

Survey question wording: “Do you support or oppose special tax rules for foreigners?” (German: In den meisten Kantonen profitieren vermögende Ausländerinnen und Ausländer von vorteilhaften Steuerkonditionen in der Form der Pauschalbesteuerung. Sind Sie für oder gegen die Pauschalbesteuerung von Ausländern?)

Policy 2 The canton has an above-average tax progressivity.

Survey question wording: “Do you prefer living in a canton with above- or below-average tax progressivity?” (German: Personen mit hohem Einkommen versteuern einen höheren Anteil des Einkommens als Personen mit tieferen Einkommen. Die Steuerprogression zwischen den Kantonen ist dabei sehr unterschiedlich. Möchten Sie persönlich in einem Kanton leben mit überdurchschnittlich starker Steuerprogression oder lieber in einem Kanton mit vergleichsweise schwacher Steuerprogression?)

Policy 3 The canton has an income tax for single persons with CHF 100k income of 12% or more.

Survey question wording: “Do you think that a single person with a gross annual income of CHF 100k should pay more or less than 12% sub-national income taxes?” (German: Finden Sie eine ledige Person mit einem Bruttoeinkommen von 100'000 Schweizer Franken sollte mehr oder weniger als 12% Einkommenssteuer an den Kanton und die Gemeinde bezahlen?)

### Immigration and Foreigners

Policy 4 The canton grants foreigners voting rights in municipal matters.

Survey question wording: “Do you think that foreigners should have the right to vote on municipal referendums?” (German: Sollten Ausländerinnen und Ausländer an Abstimmungen auf Gemeindeebene teilnehmen dürfen?)

Policy 5 The canton allows naturalization decisions to be made in town hall meetings.

Survey question wording: “Do you think that naturalizations should be decided at town hall meetings or by municipal governments?” (German: Sind Sie der Meinung das Einbürgerungsverfahren an Gemeindeversammlungen entschieden werden sollen oder sollen Fachgremien beziehungsweise der Gemeinderat über Einbürgerungen entscheiden.)

## Education and Family Policies

Policy 6 The canton grants those with children tax credits of 5,000 CHF or more.

**Survey question wording:** “Do you think that a tax credit for children of CHF 5k is too high or too low?” (German: Familien können pro Kind einen fixen Betrag vom steuerbaren Einkommen abziehen. Finden Sie 5'000 Schweizer Franken als Kinderabzug zu viel oder zu wenig?)

Policy 7 The teaching of a second foreign language starts in 5th grade or later.

**Survey question wording:** “What is the best point in time to start teaching a second foreign language – fifth grade or later?” (German: Wann finden Sie, ist der richtige Zeitpunkt eine zweite Fremdsprache zu unterrichten? Im 5. Schuljahr oder später?)

## Health-Care Policies

Policy 8 Cantonal administrations automatically contact all people eligible for health-care subsidies.

**Survey question wording:** “Do you think that public authorities should contact people who are eligible for health-care subsidies by themselves?” (German: Finden Sie es richtig, wenn die Verwaltung automatisch all jene kontaktiert, die gemäss Steuererklärung berechtigt sind Prämienverbilligungen zu erhalten?)

Policy 9 The canton grants health-care subsidies to families with two children and an annual income of CHF 90k.

**Survey question wording:** “Do you think that a family with two children and an annual income of CHF 90k should receive health-care subsidies?” (German: Finden Sie eine Familie mit zwei Kindern und einem Jahreseinkommen von 90'000 Schweizer Franken, soll Prämienverbilligung erhalten?)

Policy 10 The canton allows physicians to sell prescription medicine directly to patients.

**Survey question wording:** “Should physicians be allowed to sell medication (or only pharmacists)?” (German: Sind Sie dafür dass Ärztinnen und Ärzte in ihrer Praxis Medikamente abgeben dürfen?)

## A4 Multilevel Regression with Post-Stratification (MRP)

The following discusses in more detail the applied MRP model specifications and further investigates the elite preference measures.

### A4.1 Response

Each hierarchical response model has a number of fixed and random effects. The models used to estimate voters preferences include six random effects: sex, education, domicile type, age, canton, and region. The response models for the elite have three random effects (party, canton, region). The difference in specifications reflects the distinct weighting in the post-stratification step. In case of the voter survey, we rely on a number of individual-level variables such as age and gender that are reported in the census data. In case of the elite measure, the “population” for which we post-stratify is the political power structure of the elite. Accordingly, we use data based on party strength in the government as weights in the post-stratification of the elite measure. As fixed effects in the hierarchical model, we relied in both MRP applications on six cantonal explanatory variables: the shares of German speakers (BfS, 2012), of Roman-Catholics (BfS, 2012), of university degree holders (BfS, 2012), of left party voters in the 2011 national elections (BfS, 2013), of right party voters in the 2011 election (BfS, 2013), and GDP per capita (BAK Basel Economics).

### A4.2 Model Selection

The models were selected based on data fit determined by AIC and BIC. We rely on AIC and BIC because these measures allow us to evaluate the trade-off between data fit and model complexity. Since we are interested in predictions, over-fitting is a problem (Babyak, 2004). Relying on model quality measures that punish complexity provides a possible remedy for over-fitting. The following overview lists the variables used for the different survey questions:

- $x_{2.1}$  = % of German speakers
- $x_{2.2}$  = % of Roman Catholics
- $x_{2.3}$  = GDP per capita
- $x_{2.4}$  = % of university degree holders
- $x_{2.5}$  = % of left party voters
- $x_{2.6}$  = % of right party voters

### Selected Response Models (Voters)

1.  $y_1 \sim x_{2.1} + \alpha_{sex} + \alpha_{education} + \alpha_{age} + \alpha_{domicile} + \alpha_{canton} + \alpha_{region}$
2.  $y_2 \sim x_{2.5} + \alpha_{sex} + \alpha_{education} + \alpha_{age} + \alpha_{domicile} + \alpha_{canton} + \alpha_{region}$
3.  $y_3 \sim x_{2.4} + \alpha_{sex} + \alpha_{education} + \alpha_{age} + \alpha_{domicile} + \alpha_{canton} + \alpha_{region}$
4.  $y_4 \sim x_{2.1} + \alpha_{sex} + \alpha_{education} + \alpha_{age} + \alpha_{domicile} + \alpha_{canton} + \alpha_{region}$
5.  $y_5 \sim x_{2.5} + \alpha_{sex} + \alpha_{education} + \alpha_{age} + \alpha_{domicile} + \alpha_{canton} + \alpha_{region}$
6.  $y_6 \sim \alpha_{sex} + \alpha_{education} + \alpha_{age} + \alpha_{domicile} + \alpha_{canton} + \alpha_{region}$
7.  $y_7 \sim x_{2.1} + \alpha_{sex} + \alpha_{education} + \alpha_{age} + \alpha_{domicile} + \alpha_{canton} + \alpha_{region}$
8.  $y_8 \sim x_{2.1} + \alpha_{sex} + \alpha_{education} + \alpha_{age} + \alpha_{domicile} + \alpha_{canton} + \alpha_{region}$
9.  $y_9 \sim x_{2.4} + \alpha_{sex} + \alpha_{education} + \alpha_{age} + \alpha_{domicile} + \alpha_{canton} + \alpha_{region}$
10.  $y_{10} \sim x_{2.1} + \alpha_{sex} + \alpha_{education} + \alpha_{age} + \alpha_{domicile} + \alpha_{canton} + \alpha_{region}$

## Selected Response Models (Elite)

1.  $y_1 \sim x_{2.2} + \alpha_{party} + \alpha_{canton} + \alpha_{region}$
2.  $y_2 \sim x_{2.4} + x_{2.6} + \alpha_{party} + \alpha_{canton} + \alpha_{region}$
3.  $y_3 \sim x_{2.6} + \alpha_{party} + \alpha_{canton} + \alpha_{region}$
4.  $y_4 \sim x_{2.6} + \alpha_{party} + \alpha_{canton} + \alpha_{region}$
5.  $y_5 \sim x_{2.1} + \alpha_{party} + \alpha_{canton} + \alpha_{region}$
6.  $y_6 \sim x_{2.1} + \alpha_{party} + \alpha_{canton} + \alpha_{region}$
7.  $y_7 \sim x_{2.1} + \alpha_{party} + \alpha_{canton} + \alpha_{region}$
8.  $y_8 \sim x_{2.1} + \alpha_{party} + \alpha_{canton} + \alpha_{region}$
9.  $y_9 \sim x_{2.2} + x_{2.4} + \alpha_{party} + \alpha_{canton} + \alpha_{region}$
10.  $y_{10} \sim x_{2.1} + x_{2.5}\alpha_{party} + \alpha_{canton} + \alpha_{region}$

### A4.3 Exogeneity of Elite Responses

One concern is that politicians do not answer the questions sincerely, but give answers that they believe are in line with their constituency. If that were the case, the validity of the model estimates would be questioned. We show this is most likely not the case by comparing the estimates with respect to their constituency and their party. The analysis presented below investigates the extent to which politicians’ responses are a function of their party affiliation and the position of their constituency. [Table A4](#) reports the results for policy question 1 (the results are robust for all 10 policy questions).

	Model 1	Model 2	Model 3
Constant	0.86*** (0.03)	0.30*** (0.10)	0.60*** (0.07)
FE PARTIES	✓	X	✓
FE CANTONS	X	✓	✓
R <sup>2</sup>	0.63	0.10	0.67
BIC	257.7	740.3	357.8

\*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

Table A4: Exogeneity of Elite Preferences

If politicians’ answers were to a certain degree endogenous, they should vary by constituency and not by party. However, the model with the best BIC value only includes the party indicators (Model 1), which shows that party affiliation is a very strong predictor of a cantonal politician’s policy preferences. This result is in line with more recent research showing that politicians might be ill-equipped to estimate their constituency’s preferences. [Broockman and Skovron \(2015, 1\)](#) document with a large survey of 2,000 legislative candidates that “actual district opinion explains only a modest share of the variation in politicians’ perceptions of their districts’ views.” Of relevance in that regard may be that the elite survey was carried out anonymously, which should increase the chances that the politicians state their true preferences.



## A5 The Direct Democracy Measure

The presented empirical analyses rely on a widely used direct democracy index measuring how difficult it is to use direct democratic rights (Stutzer, 1999; Frey and Stutzer, 2000). The index is in line with the theoretical models, as it approximates quite precisely the costs of citizens to use direct democratic instruments (for an alternative measure, see Appendix A6). Figure 5 shows the substantial variation in direct democracy across Swiss cantons, ranging from a value of 1.75 (low direct democracy) in the canton of Geneva to 5.5 in the canton of Glarus. The map also shows that direct democracy is more comprehensive in the German speaking part of Switzerland.

For putting the findings into context, an important question is how direct democratic the Swiss cantons are in comparative perspective. If we apply the calculation of the direct democracy index to California, we calculate a rather low value of 2.33. However, the applied index is tailored to the Swiss version of direct democracy, which biases this comparison. For example, California is much more direct democratic in the sense that the government is neutral in the campaign, political representatives have to implement successful initiatives, and citizens can ask for recall elections. In the Swiss case, however, the governments and parliaments are quite heavily involved in the initiative process: they can draft counterproposals, are active in campaigns, and finally, in case of successful initiatives, they interact with the authors of the initiative for the elaboration of implementation laws.

In sum, taking the comparison (and the caveats) into account, the variation in direct democracy among Swiss cantons is very substantial and it spans from limited direct democracy to very comprehensive participatory rights.

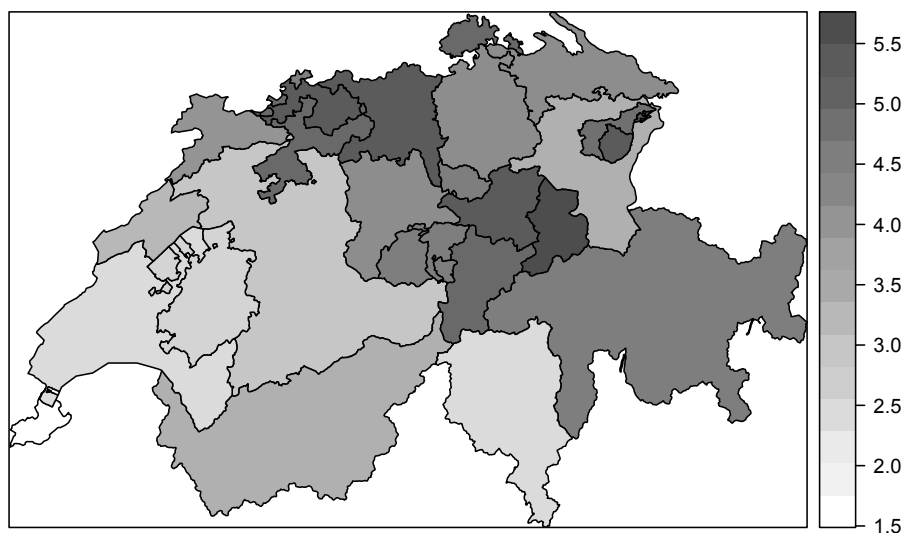


Figure 5: Direct Democracy Index Values for Swiss Cantons

## A6 Robustness Checks

As an alternative to the [Stutzer \(1999\)](#) index of direct democracy, we check the robustness of the results with new data on direct democratic institutions in the Swiss cantons from [Vatter et al. \(2010\)](#). The findings presented in [Table A6](#) are robust.

	Model 1	Model 2	Model 3	Model 4
DIRECT DEMOCRACY (ALT.)	0.79** (0.37)		0.51 (0.42)	-1.05 (0.76)
GOVERNMENT-VOTER PREFERENCE DEVIATION		-2.57** (1.00)	-10.34*** (3.65)	-11.49*** (3.86)
DD (ALT.) x PREFERENCE DEVIATION			10.75** (4.84)	13.35*** (5.07)
VOTER OPINION CLARITY				3.72** (1.55)
GOVERNMENT OPINION CLARITY				-2.32** (1.11)
GDP (PER CAPITA IN 100K)				0.04 (0.59)
ELECTORAL THRESHOLD				-0.01 (0.01)
% GERMAN SPEAKERS				1.22** (0.58)
LEGISLATIVE POWER				0.01 (1.12)
POPULATION SIZE (IN 100K)				-0.02 (0.04)
SVP VOTE SHARE				-0.01 (0.02)
DD IN USE (PAST 3 YEARS)				-0.27 (0.27)
CONSTANT	-0.50 (0.31)	0.14 (0.19)	-0.15 (0.32)	0.15 (1.15)
BIC	360.59	359.21	359.12	392.25
$\ell\ell$	-169.17	-168.48	-162.88	-154.42
<i>N</i>	260	260	260	260
Groups: Cantons	26	26	26	26
Groups: Policies	10	10	10	10
Variance: Canton	0.03	0.07	0.04	0.02
Variance: Policy	0.33	0.25	0.26	0.29

*CPC*: Correctly predicted cases; baseline is 51%. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

Table A6: Robustness Check With Alternative Direct Democracy Measure.

As an additional robustness check, we use an alternative measure of the citizen-government preference deviation. In the main analysis, we set the government-voter preference gap equal to 0 when the government and voters share the same majority opinion. The alternative preference deviation measure also differentiates between the sizes of the gaps in that case. Concretely, we estimate the absolute difference of the preference gap and, when the the government and voters hold opposing majority views, we add the value of 50 to the absolute deviation. With this coding, we derive a gap variable that theoretically ranges between 0 and 150 and differentiates between congruent (0 to 50) and incongruent (51 to 150) majority preferences. The smaller the values of the variable, the closer are the positions of the government and voters. The presented findings are robust.

	Model 1	Model 2	Model 3	Model 4
DIRECT DEMOCRACY	0.19** (0.09)		0.09 (0.11)	-0.19 (0.19)
GOVERNMENT-VOTER PREFERENCE DEVIATION (ALT.)		-0.93*** (0.36)	-3.65** (1.52)	-4.04** (1.58)
DD X PREFERENCE DEVIATION (ALT.)			0.60* (0.33)	0.73** (0.34)
VOTER OPINION CLARITY				4.01** (1.57)
GOVERNMENT OPINION CLARITY				-2.55** (1.09)
GDP (PPC IN 100K)				0.07 (0.60)
ELECTORAL THRESHOLD				-0.01 (0.01)
% GERMAN SPEAKERS				0.99* (0.59)
LEGISLATIVE POWER				0.10 (1.13)
POPULATION SIZE (IN 100K)				-0.01 (0.04)
SVP VOTE SHARE				-0.00 (0.02)
DD USED (PAST 3 YEARS)				-0.26 (0.28)
CONSTANT	-0.81* (0.41)	0.24 (0.21)	-0.12 (0.51)	0.31 (1.31)
BIC	360.19	358.99	360.35	394.23
$\ell\ell$	-168.97	-168.37	-163.49	-155.41
-----				
$N$	260	260	260	260
Groups: Cantons	26	26	26	26
Groups: Policies	10	10	10	10
-----				
Variance: Canton	0.03	0.06	0.04	0.03
Variance: Policy	0.33	0.26	0.27	0.33

*CPC*: Correctly predicted cases; baseline is 51%. \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

Table A6: Robustness Check With Alternative Preference Deviation Measure.

## A7 Policy Dimensionality Analysis

The finding that direct democracy is most effective when the elite-voter gap is largest raises the question of when we would expect that gap to be large and thus direct democracy to be an effective democratic corrective. While representation scholars emphasize that elections are powerful for minimizing the elite-voter preference gap (Schumpeter, 1942; Mansbridge, 2009), the advantage of direct democracy is that referendums and initiatives allow it to unbundle policy issues (Kriesi, 2005; Besley and Coate, 2008).

The analysis of this study, the insights of the representation literature, and the unbundling argument all suggest that the effectiveness of direct democracy on policy congruence depends on the dimensionality of the political conflict structure. Accordingly, we should expect that elections are powerful in democratically accommodating the main political conflict dimension and that direct democratic institutions are particularly effective in aligning policies on second-ordered dimensions with the majority will of the electorate. If that is correct, both the citizen-elite preference gap and the effect of direct democracy on policy congruence should increase as we move away from questions of the dominant (first) to other policy dimensions. The first dimension usually covers classic left-right distributional questions, while the second dimensions include cultural, identity, and residual issues (Kriesi et al., 2008; Kitschelt and Rehm, 2014).

In the following, we present an empirical policy dimensionality investigation. Before we present the analysis, however, we would like to emphasize that our data does not allow us to analyze the dimensionality of the political conflict structure for each canton separately. To conduct the comparative analysis, we thus have to make simplifying assumptions based on the findings of the literature. We assume that there are two dimensions plus a residual category. Concretely, we code the questions on the tax rate, the progression, and the subsidies for health-care as distinctive left-right distributional policies; we group all the three policies targeted at foreigners in the second dimension category; and the remaining policies (education, family, and health) form the a residual category.

Figure 6 and Table A7 show how the absolute preference gaps between the elite and the citizens vary in the three categories and whether the effect of direct democracy varies across the policy dimensions. The box plots of Figure 6 show that the elite-voter gap is smallest in the policies of the first dimension. The estimates in Table A7 show a positive interaction effect between the policy dimension variable and direct democracy, which suggests that the effect of direct democracy is driven by policies not covered by the first dimension. The findings support the argument that the effect of direct democracy depends on the dimensionality of the political conflict structure.

However, we would like to emphasize that this analysis comes with limitations. Thus, the findings should be interpreted carefully. Most importantly, we assume that the political conflict structure is constant across cantons—despite the heterogeneity among the cantonal polities. Some cantons are dominated by one or two parties, while we find in others a more fragmented party system with many parties and consociational multiparty governments. Our data do not

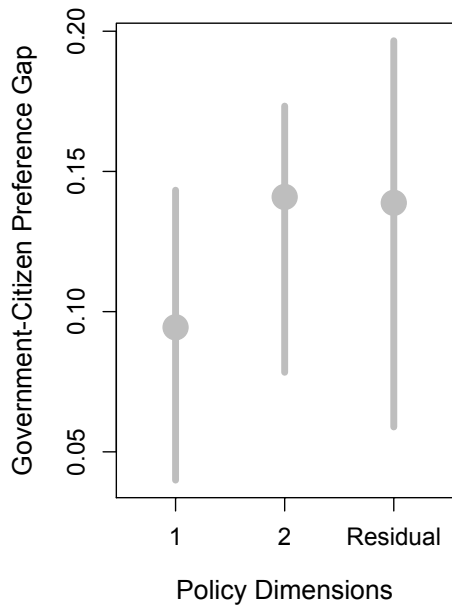


Figure 6: Preference Gap of Policy Dimensions (1st Quartile, Mean, 3rd Quartile)

	Model 1	Model 2
DIRECT DEMOCRACY	0.22** (0.09)	-0.23 (0.21)
GOV.-VOT. PREF. GAP	-2.95*** (0.98)	-3.13*** (0.99)
POLICY DIMENSION	0.32* (0.18)	-0.58 (0.43)
DD x POLICY DIM.		0.21** (0.09)
CONSTANT	-1.42*** (0.55)	0.45 (0.99)
BIC	360.52	361.18
$\ell\ell$	-163.58	-161.13
-----		
$N$	260	260
Groups: Cantons	26	26
Groups: Policies	10	10
-----		
Variance: Canton	0.04	0.04
Variance: Policies	0.15	0.16

\*\*\* $p < 0.01$ , \*\* $p < 0.05$ , \* $p < 0.1$

Table A7: The Direct Democracy Effect on Policy Congruence Conditional on Policy Dimensions

allow us to make a fine-grained dimensionality analysis for each canton. In that sense, our findings provide a broader macro picture on the question of how direct democracy interacts with the political conflict structure in Swiss cantons. Given the limitations, the findings of the dimensionality analysis are suggestive. More empirical research is needed to further disentangle the link between political conflict structure, policy dimensionality, and the effectiveness of direct democracy on policy congruence.

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