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## ***The beauty in a beast: Minimising the effects of diverse recording quality on vowel formant measurements in sociophonetic real-time studies***

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### **ABSTRACT**

Sociophonetic real-time studies of vowel variation and change rely on acoustic analyses of sound recordings made at different times, often using different equipment and data collection procedures. The circumstances of a recording are known to affect formant tracking and may therefore compromise the validity of conclusions about sound changes made on the basis of real-time data. In this paper, a traditional F1/F2-analysis using linear predictive coding (LPC) was applied to the vowels /i ʌ a/ extracted from spontaneous speech corpora of Glaswegian vernacular, that were recorded in the 1970s and 2000s. We assessed the technical quality of each recording, concentrating on the average levels of noise and the properties of spectral balance, and showed that the corpus comprised of mixed quality data. A series of acoustic vowel analyses subsequently unveiled that formant measurements using LPC were sensitive to the technical specification of a recording, with variable magnitudes of the effects for vowels of different qualities. We evaluated the performance of three commonly used formant normalisation procedures (Lobanov, Nearey and Watt-Fabricius) as well as normalisations by a distance ratio metric and statistical estimation, and compared these results to raw Bark-scaled formant data, showing that some of the approaches could ameliorate the impact of technical issues better than the others. We discuss the implications of these results for sociophonetic research that aims to minimise extraneous influences on recorded speech data while unveiling gradual, potentially small-scale sound changes across decades.

*Keywords: Real-time corpus; Formants; Formant normalisation; Noise; SNR; Spectral tilt; Sociophonetic 'gold standard'*

## 1. INTRODUCTION

### *1.1. On the issue of comparability in sociolinguistic data*

Since its origins in the early 1960s, variationist sociolinguistics has been concerned with the methodological rigour of its quantitative enquiry. In the foreground of the early discussions were the issues primarily involving the data collection, such as the “Observer’s Paradox”, style shifts and sampling strategies (Labov 1972, Cukor-Avila, 2000). Subsequent studies have further unveiled the multitude of the potential sources of influences in sociolinguistic data, which include (and are not limited to) familiarity between the participant and the interviewer, presence of additional peers during the interview, the experience and elicitation strategies of the interviewer as well as the quantitative approaches to analysing the data (Labov 1972; Milroy 1987; Milroy and Gordon 2008; Llamas et al 2006; Tagliamonte 2006; see Tillery and Bailey (2003) for a critical overview). All of these factors may not only influence the observed results, thus misleading generalisations about the patterns of variation and change, but also reduce comparability of the results across different studies of the same sociolinguistic phenomena, undermining the core principles of methodologically sound research, reliability and intersubjectivity.

Ultimately, sociolinguistic research aims to combine natural (or at least naturalistic) data which preserves the social identity (Scobbie and Stuart-Smith 2012) with a rigorous amelioration of any extraneous influences that can influence the data patterns. In their critical paper, Tillery and Bailey (2003) suggested that this standard can only be achieved through a solid understanding of the sources and the magnitudes of possible extraneous influences on sociolinguistic data patterns, and regretted the current lack of such understanding, calling for more research in this methodologically highly relevant area.

The present study aims to contribute to this endeavour, and is concerned with the potential influences of technical specifications of recordings on the vowel formant measurements taken from them. Vowel formants are the core acoustic correlates of vowel quality typically obtained in sociophonetics (but see Harrington, Kleber and Reubold (2013) for an alternative set of acoustic measures), and have been scrutinised in many studies of sound variation and change (e.g. Fought 1999; Harrington, Cox and Evans 1997; Labov 1994; Labov, Ash and Boberg 2006; Maclagan et al. 2009; Mesthrie 2010). In an apparent-time setting, much care has traditionally been taken to account for the formant differences arising from speaker physiology, relating primarily to the age and the vocal tract size (e.g. Linvillea and Rens 2001), and to distinguish these physiological influences from the sociolinguistically relevant patterns produced by speakers of different ages and sexes (e.g. Labov, Ash and Boberg 2006). Numerous techniques have been developed, tested and compared in order to achieve the normalisation for speaker physiology while preserving the social indexicality of their speech (e.g. Adank, Smith and van Hout 2004; Clopper 2009; Watt and Fabricius 2002; see Flynn (2011) for an overview). We will discuss the most commonly used approaches in Sect. 3.3 below.

In contrast to this long-standing methodological debate characteristic of apparent-time studies, real-time studies of sound variation and change have rarely problematized potential issues involved in formant measurements of vowels. Trend studies with real-time data (recorded with different samples of individuals from the same community at different points in time) are unanimously recognised as a particularly insightful and reliable methodological setting for studying language change at a community level (e.g. Labov 1994; Sankoff and Blondeau 2007; Trudgill 1988), primarily because they eliminate effects related to speaker age, such as age grading (Wagner 2012). However, real-time studies frequently rely on acoustic analyses of recordings of speech made using different equipment with variable technical specifications and following different recording procedures. To date, still little is known about the sources, types and magnitudes of technical influences on the formant data. In the next section, we will give an overview of the currently established effects, and hypothesise how they might play out in a real-time study of sound variation and change.

### ***1.2. Technical influences on formant measurements***

Not many studies have addressed the question of whether, and how, formant values (extracted using the traditional method of LPC) might be influenced by the equipment and set-up of a recording and its resulting technical specifications. A series of studies have been conducted in the context of forensic speaker identification (e.g. Byrne and Foulkes 2004; Künzel 2001); and only a few, mostly preliminary investigations have recently pointed out that technical issues of a recording may obscure the patterns of variation and change in sociophonetics, too (De Decker and Nycz 2011; De Decker 2016; Hansen and Pharao 2006; in progress).

In terms of the recording equipment and set-up, several features have been identified to leave an imprint in the vowel spectrum and to impact on the measured formant values. First of all, the band-pass filtering due to the transmission by phone lines (both mobile and landline) is known to interfere with the calculation of the formants (Byrne and Foulkes 2004; Künzel 2001). Harmonics that lie below the lower cut-off boundary (approximately 300 Hz) and above the upper boundary (approximately 3.2 kHz in mobile phones and 3.5 kHz in landline transmissions) are most affected, since their weighting in the calculation of the formant frequencies is decreased. This usually leads to artificially high frequencies of F1 (particularly in high vowels whose F1 is much stronger affected than the relatively high F1 of low vowels). However, even F2 whose frequencies fall within the transmitted range shows some technically introduced artefacts. In comparison to the values obtained from a recording made simultaneously with a studio microphone, F2 of high vowels tends to measure lower values in mobile recordings (Byrne and Foulkes 2004), though the effect tends to be smaller and has not been consistently documented in other phone transmissions (Künzel 2001). The exact magnitudes of these technically introduced effects also seem to vary substantially across different studies and types of phone transmissions, and range between 14 and as high as 60 percent of the original frequency (Byrne and Foulkes 2004; Künzel 2001).

Similar to the effects of band-pass filtering for a cost-effective phone transmission, compression algorithms used for a space-effective storage of video and digital audio recordings (as e.g. available on the internet) have been shown to influence spectral properties of speech recordings (De Decker and Nycz 2011; Rozborski 2007; van Son 2005). F1 seems to be affected across the board, measuring higher values after a compression, while the impact on F2 is rather mediated by vowel quality, raising F2 in high vowels but lowering it in low vowels (De Decker and Nycz 2011). Again, the magnitude of these effects varies across studies and compression methods, ranging from negligible ( $\leq 3\%$ , van Son 2005) to quite substantial (De Decker and Nycz 2011), with higher compression rates leading to a more significant distortion of the original recording (Rozborski 2007). Although mobile devices admittedly introduce numerical artefacts in the formant values during the transmission (cf. Byrne and Foulkes 2004), De Decker and Nycz (2011:54) argue that recordings made with some portable devices of the same manufacturer (here, Macbook Pro and iPhone) produce comparable measurements, and maintain an overall shape and size of the vowel space in comparison to uncompressed recordings (at least as far as F1 and F2 are concerned), thus lending themselves to a sociolinguistic investigation better than others (e.g. Mino-derived formats commonly used by YouTube).

Apart from the influence the format of a recording can have on its spectra and formant measurements taken using LPC, somewhat less obvious factors, such as ambient noise, room acoustics, microphone make and placement during the recording session, have also been shown to leave their spectral imprints and interfere with formant measurements (De Decker 2016; Hansen and Pharao 2006, in progress; Plichta 2004; Švec and Granqvist 2010). The quality of the recordings not controlled for such influencing factors will likely vary with respect to at least two technical specifications (cf. Švec and Granqvist 2010): (1) the levels of noise, typically measured by the signal-to-noise ratio, SNR (see 2.3) and (2) spectral balance (or tilt), reflected in the distribution of the intensity across lower vs. higher harmonics of the spectrum (see 2.3 for more detail).

It is well known that high levels of background noise reduce intelligibility of speech (e.g. Pollack and Pickett 1958), but even recordings made in relatively quiet surroundings can differ with respect to their SNR. For example, hiss (or low-level white noise) can originate from analogue electronics, ground hum and buzz from improperly grounded systems: the fundamental of 50 or 60 Hz and their harmonics will be distinguishable in the recording spectrum (Corley 2010). An increased distance of the microphone from the sound sources can also decrease SNR, making the room reverberation and noises more prominent in a sound recording (Corley 2010:57). Omnidirectional microphones usually pick up more background noise than directional ones, with the small-tip versions producing particularly noisy recordings (Švec and Granqvist 2010). In such increased noise levels (reflected in lower SNR, see 2.3), formants often appear very faint or have larger bandwidths and are therefore less clearly defined (Plichta 2004); Plichta strongly advises against using such

recordings for speech research. De Decker (2016), however, shows that not all types of background noise have an equally damaging effect on the accuracy of formant estimation. High levels of white noise (i.e. a signal with an equal amplitude at all frequencies of its wide-band spectrum) exerts a particularly strong impact on the formant estimation while a 60-Hz hum and even speech babble only have a subtle effect, if any at all.

Moving the microphone closer to the sound source – i.e. speaker mouth – may solve the problem of the background noise in some cases (cf. recommendation (3) in De Decker 2016:99), however, a too close placement is likely to cause the so-called proximity or bass effect, referring to a strong boost of lower frequencies in a spectrum (Corley 2010; Švec and Granqvist 2010). All directional microphones are known to boost low frequencies when working close, while omnidirectional microphones are relatively free of the effect (Corley 2010; Švec and Granqvist 2010). Moreover, each microphone has its individual frequency response, i.e. the intensity levels of the recording over the operating range of the frequencies. In some microphones, a flat frequency response cannot be guaranteed at distances other than 30 cm away from the sound sources while others always amplify the frequencies between 3 and 10 kHz regardless of the distance (Švec and Granqvist 2010). Brixen (1996; 1998) shows that differences in the microphone placement result in different power spectra, particularly affecting lavalier and headband microphones when they are placed very close to the sound source. Such differences in power spectra have been further documented to impact upon LPC-based formant measurements, and to result in partly substantial discrepancies between formant values extracted from these recordings (Brixen 2011; Hansen and Pharao 2006, in progress): once again, F1 seems to be more affected than F2, and shows discrepancies of up to 5 semitones while F2 measurements deviate from each other in the region of 2-3 semitones when taken from recordings with different spectral specifications. Moreover, the differences in spectra that lead to deviating formant values may result not only from the microphone placement, but also from a specific frequency response each microphone has as part of its technical specification and often depends on the microphone-to-sound distance (Švec and Granqvist 2010 for an overview). Importantly, Hansen and Pharao (in progress) highlight that these differences in the LPC-based formant measurements cannot be straightforwardly attributed to the microphone placement alone, but interact in complex ways with several other external factors, including the particular vowel quality being measured and the room acoustics where the recording took place.

To summarise results of the previous research discussed in this section, there is sufficient evidence to demonstrate that LPC-based formant measurements are sensitive to the technical quality of recordings. Two main recording features seem to play a major role: (1) the SNR levels which can vary depending on the presence of extraneous noise, the microphone type and the exact distance between the microphone and the mouth of the speaker, and (2) the distribution of spectral information across the frequency range of 0-5 kHz varies in the recordings which depends on a variety of factors related mainly to the recording equipment and set-up. In the resulting measurements, the frequency of F1 is

known to be particularly affected by these issues, although F2 also shows some technically induced artefacts, even if of a smaller magnitude.

The primary interest of the studies reviewed in this section is rather technical (with implications for forensic or sociophonetic research, e.g. De Decker 2016; De Decker and Nycz 2011; Rozborski 2007). Most of them compared recordings made simultaneously with variable equipment, and estimated the resulting differences in formant frequencies those recordings measured (Byrne and Foulkes 2004; De Decker and Nycz 2011; Hansen and Pharao 2006, in progress; Künzel 2001; Plichta 2004; van Son 2005). While all of the factors identified above may potentially play out in sociophonetic real-time corpora and contribute to the corpus diversity in terms of technical quality, the ultimate goal of sociolinguistic research is to unveil gradual, potentially small-scale sound variations and changes across decades. For doing so, researchers need to be aware of the sources, the magnitudes and directions of such technical influences and enabled to keep their imprint on recorded speech data minimal (see 1.1). The present paper aims to tackle this challenge, or at least evaluate different strategies to approaching this methodological issue. To our knowledge, no previous study has identified and systematically investigated the relationship between such properties of the spectrum as variable SNR levels and spectral tilts on the one hand, and the LPC-based formant values on the other. These two features of the spectrum may result from various sources over which sociophoneticians working with real-time corpora may not have control, but the knowledge of the existence of these two easily assessable (see 2.3) spectral properties may well enable them to apply some post-hoc procedures that will eliminate the technically introduced artefacts in formant values described above. We can test and evaluate such possible post-hoc procedures using a phonetic case study where the presence of change and its direction have been demonstrated by independent research - this will allow us to see if we can minimise the potential effects of technical diversity in a real-time corpus, while still observing patterns of variation and change that converge with the previous findings. A perfect candidate for such a case study was deemed to be the variable /u/ in Scottish English.

### ***1.3. /u/ in Scottish English***

As a case study into sound change, the high back round vowel /u/ has recently attracted an extensive experimental scrutiny in many varieties of English. A change in progress towards a more fronted, near-central variant [ɯ] has been demonstrated for many varieties of English, including Standard Southern British English (e.g. Hawkins and Midgley 2005; Harrington 2007; McDougall and Nolan 2007), American English spoken by mainstream as well as minority speakers (Labov, Ash and Boberg 2006; Fought 1999), Australian, New Zealand and South African varieties (e.g. Harrington, Cox and Evans 1997; Maclagan et al 2009; Mesthrie 2010). While this change appears to have been taking place in Anglo-English over the last fifty years (e.g. Harrington et al 2011), the situation seems to be more complex north of the border in Scotland. A fronted production of this vowel has been a

long-established diagnostic trait in Scottish varieties of English, most notably in the urban Scots vernacular spoken in Glasgow (e.g. Wells 1982: 402; Stuart-Smith 2004: 58–59).<sup>1</sup>

Auditorily, Scottish English /u/ has been reported to be close to the central [ʊ] for a long time; vernacular Scots is reported to be even fronted in contrast to the backer [u] of educated Scottish Standard English (McAllister 1938; Macaulay 1977; Johnston and Speitel 1983). Contemporary articulatory-phonetic analysis suggests that /u/ is not only front but also low, and specifically that the tongue position is as front as that of front vowels /i ɪ e ε/, and lower than /ɪ/ and /ε/ (Scobbie 2011; Scobbie, Lawson and Stuart-Smith 2012). These findings may reflect a real-time process of fronting and lowering, i.e. we can expect to detect a change in the acoustic realisation of this vowel in real-time over the past 40 years. Recent real-time acoustic-phonetic data from Glasgow, comparing speakers from 1916 with those recorded in the 1970s to 2000s, of different ages, suggests that /u/ has lowered over the 20th century (Stuart-Smith et al 2016). Taken together, these findings suggest a real-time process of lowering, and possibly further fronting, i.e. we can expect to detect a change in the acoustic realisation of this vowel in real-time over the past 40 years.

#### **1.4. Goals of the present study**

The main goal of the study was to examine whether or not differences in the technical quality of recordings, which seemed likely to intersperse a sociolinguistic real-time corpus (SNR, spectral balance), may have an impact on the values of F1 and/or F2 calculated using a standard LPC-algorithm. We addressed this question using the example of a subset from an electronic real-time corpus of Glaswegian vernacular speech which comprises of diverse recordings made at different points in time, by different people and for different purposes (including sociolinguistic and oral history interviews as well as free conversations, see 2.1; Rathcke and Stuart-Smith 2015; Stuart-Smith et al. 2015).<sup>2</sup> We chose /u/ as a case study into disentangling the technical effects from the sound change, given there exists some reliable external evidence for this vowel in both historical and modern-day Scottish English data (McAllister 1938; Macaulay 1977; Johnston and Speitel 1983; Scobbie 2011; Scobbie, Lawson and Stuart-Smith 2012; Stuart-Smith et al 2016).

If, as we may expect (see 1.2), a technical influence can indeed be attested in formant measurements taken from sociolinguistic real-time recordings, we would further want to know if different vowel qualities were affected in similar ways, and the extent to which they were affected. For the purposes of this investigation, we chose the high front vowel /i/ and the low central /a/ as reference vowels. Apart from being the corner vowels of the Scottish system (Scobbie 2011; Scobbie, Lawson and Stuart-Smith 2012), neither /i/ nor /a/ could have been expected to show sound changes in Scottish English (see 3.5 for

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<sup>1</sup> It should be noted that Glaswegian has a typical vowel systems of Scottish English characterised by a FOOT–GOOSE merger (e.g. Ferragne and Pellegrino 2010; Stuart-Smith 2004).

<sup>2</sup> For more information on this real-time project, visit <http://soundsofthecity.arts.gla.ac.uk/index.html>



further discussion). In contrast, our target vowel – which we will henceforth refer to as the high rounded central /ɤ/ - may well be in focus of potential sound changes (see 1.3).

Finally, following from the above findings, we aimed at establishing a procedure to minimise, or even eliminate, any effect of technical quality differences on F1/F2 measurements in order to allow for a methodologically sound inference of vowel changes in valuable real-time corpora which are generally recognised to offer an insightful and reliable tool for studying language change at a community level (e.g. Labov 1994; Sankoff and Blondeau 2007; Trudgill 1988).

## 2. DATABASE

### 2.1. Real-time corpus of Glaswegian vernacular speech

The real-time data to be discussed in this paper span four decades of Glaswegian vernacular speech, from early 1970s to late 2000s. They draw upon common types of speech recordings which were deemed representative of the data available for sociolinguistic real-time analyses at the present day (i.e. sociolinguistic or oral history interviews, free conversations). Recordings of this type are typically held at national or local libraries, sound archives or in private collections. This paper deals with data from 24 male speakers of the working class background. The sample consists of three age groups: **Young** (12-17 y.o.), **Middle-aged** (40-55 y.o.) and **Old** (67-90 y.o.), with four speakers per group. The speakers were recorded either in the 1970s or in the 2000s.

A large part of the 1970s-subcorpus consists of sociolinguistic interviews carried out by Ronald Macaulay in 1973 (Macaulay 1977). Teenagers (*70-Y, m1-m4*) and middle-aged speakers (*70-M, m1-m3*) were interviewed in quiet rooms using a lavalier microphone *Uher M822* which was placed somewhere on the speaker (usually on the lapel). These recordings are held at the University of Edinburgh and were digitised at their Sound Archives. The remaining speaker of the middle-aged sample (*70-M, m4*) was interviewed by William Labov in the early 1970s. The interview took place at the interviewee's home, in the East End of Glasgow. Details of the recording equipment are not known. The digitised sound file was kindly provided by the Linguistics Department of the University of Pennsylvania. For the older speaker group, most of our recordings (*70-O, m1-m3*) were collected between 1970 and 1973 for the project entitled '*Family Life and Work Experience before 1918*' as a part of the first national oral history survey in the UK (Thompson 1975/1992). The master tapes of these recordings are held at the British Library in London and were digitised by their media services. They were supplemented by an interview about the history of the cinema and film theatres in Glasgow, which was recorded in the early 1970s at the People's Palace by the curator of the museum Elspeth King. The cassette recording is held at the Glasgow Museums Resource Centre and was digitised by the first author using Marantz CP 430 recorder.

The main resource for speech data from the 2000s was the Media Project Corpus recorded in Glasgow in 2003, along with socio-economic and attitudinal data, to investigate the

impact of media on linguistic properties of the Glaswegian vernacular (Stuart-Smith 2006; Stuart-Smith et al. 2013). Free conversations between self-selected pairs of interlocutors included teenagers (*00-Y, m1-m4*), who were recorded in quiet rooms at schools, and adult speakers (*00-M, m1-m04*), who talked to each other in Glasgow West-End pubs with varying degrees of background noise interspersing the conversation. For these recordings, a *Sony ECM CS10* lavalier microphone was placed on the speaker's chest. The data for the older speaker group was taken from two series of oral history interviews, *The M-74* collection (*00-O, m1*, recorded in 2008) and *The Dock Workers* collection (*00-O, m4, m5, m6*, recorded in 2010), both conducted by an oral historian from the Scottish Oral History Archives at the University of Strathclyde. *Audio-technica PRO70* lapel microphones were used and placed on the speaker's upper chest. The interviews took place at interviewee's homes, in quiet surroundings. Digitised copies of the interviews were made available for phonetic research by the Archives.

We note that the recordings of the corpus were digitised at varied sampling rates, using diverse equipment. Little is known about how the digitisation of old reel-to-reel tapes (as in Macaulay's recordings) or cassettes (as in all other 1970s-recordings) may affect acoustic properties of the spectrum. A discussion of such issues is beyond the scope of this paper. For the purposes of the current investigation, all recordings were downsampled to 10 kHz for the formant analyses undertaken in Praat (see below for more detail).

The recordings might differ with respect to their stylistic setting due to the field method employed for their collection (see Gregersen and Barner-Ramussen (2011) for an overview). 2000s-dialogs of young and middle-aged men are quite possibly a closer approximation of vernacular speech since they had the advantage of a familiar audience as compared to the sociolinguistic and oral history interviews available for the 1970s and older speakers (Bell 1984). To this end, *00-M* and *00-Y* data obtained from more casual conversations can be expected to show more target undershoot, i.e. more centralised realisations of all vowels (Moon and Lindbloom 1994; Picheny, Durlach and Braida 1986). Also, a stylistic shift towards the local standard variety may occur in more formal interview settings. In the context of this investigation, this would mean a slight retraction of the target vowel /ʌ/ (cf. Stuart-Smith 1999) and is more likely to occur in the sociolinguistic and oral history interviews (all 70s and *00-O* groups) than in the spontaneous conversations (*00-Y* and *00-M* groups). However, the situation of being recorded while speaking into a microphone might have created comparable levels of awareness and attention drawn to speech production and therefore resulted in negligible phonetic differences due to style shifting (cf. Labov 1994: 157-158). In any case, even if we consider stylistic differences across the dataset as being minimal, we are still left with conspicuous differences in technical quality of the 1970s and 2000s-recordings, to be illustrated below (see 2.3).

## **2.2. Data preparation**

All corpus recordings were first transcribed orthographically by native speakers of Scottish English. Dysfluencies, overlaps, laughed or sung speech and other features worth

considering in subsequent analyses were captured at multiple layers in Transcriber software (Barras et al. 2001).

For the chosen 24 speakers, we extracted all words containing lexically stressed and phrasally prominent /i ʌ a/ vowels (except those preceding a postvocalic /r/), totalling N=3610. To insure consistency across the dataset, a protocol of segmentation and labelling was developed to guide the data preparation by two fully trained phoneticians (one of them the first author). EMU-software (Cassidy and Harrington 2001, Winkelmann 2015) was used to create a hierarchically and sequentially organised speech database for acoustic analyses. All acoustic measurements reflecting recording quality reported below were taken based on the DFT-spectra created in EMU (Harrington 2010, see 2.3) while formant values were calculated using Praat (Boersma and Weenink 2013, see 3.1). Subsequent data processing was conducted in R (version 2.15.1). Statistical analyses were also run in R (version 3.1.0).

### **2.3. Differences in recording quality across the selected dataset**

To our knowledge, most of the recordings in our corpus were made using lavalier omnidirectional microphones. However, they were of different makes, and we know nothing about their frequency response. As far as we are aware, the microphone placement was not controlled for, neither were sources of background noise or acoustic properties at the respective places where recordings took place. To illustrate resulting differences in sound quality, spectrograms of six recording samples showing the frequent word *good* (chosen as it contains the target vowel /ʌ/) are compared in Figure 1.

The two main issues related to the technical quality of recordings discussed in 1.2 can be confirmed for these data: (1) the SNR levels are highly variable, potentially reflecting both different levels of extraneous noise and varied distance between the microphone and the mouth of the speaker, and (2) the distribution of the spectral information across the frequency range of 0-5 kHz varies in the recordings, potentially reflecting differences in the make and the placement of the microphones used and the acoustics of rooms where recordings took place. Similar to the observations made in previous research (see 1.2), we find that poor SNR can make formants appear very faint (cf. *00-0, m3*) or have larger bandwidths and therefore be less clearly defined (cf. *00-M, m1*).

Weak acoustic information in the high frequency range (above 4 kHz in *70-0, m3* or even above 2.5 Hz in *70-M, m4* and *70-Y, m2*) accompanied by higher intensity of F1 seems to be a particular feature of some 1970s-samples. The distribution of acoustic energy across the vowel spectrum has a much steeper negative spectral slope as compared to a more balanced spectrum as in e.g. the *00-Y* sub-corpus, making F2 less well defined in the vowel spectrum of the *70-Y* speakers. This difference between recordings made in the 1970s and the 2000s is illustrated in Figure 2. It seems unlikely that the proximity effect alone might have caused this spectral imbalance (Švec and Granqvist 2010), given that the first spectral peak is not substantially higher in the *70-Y, m2* than in the *00-Y, m3* example (Figure 2). However, the second and the third peaks appear much lower in the earlier recording. We

considered the possibility of these slope differences being due to differences in the voice quality (e.g. Hillenbrand, Cleveland and Erickson 1994), but our perceptual analyses of the speakers' voices did not confirm this potential explanation. We will return to this issue in the discussion (see 5.4).

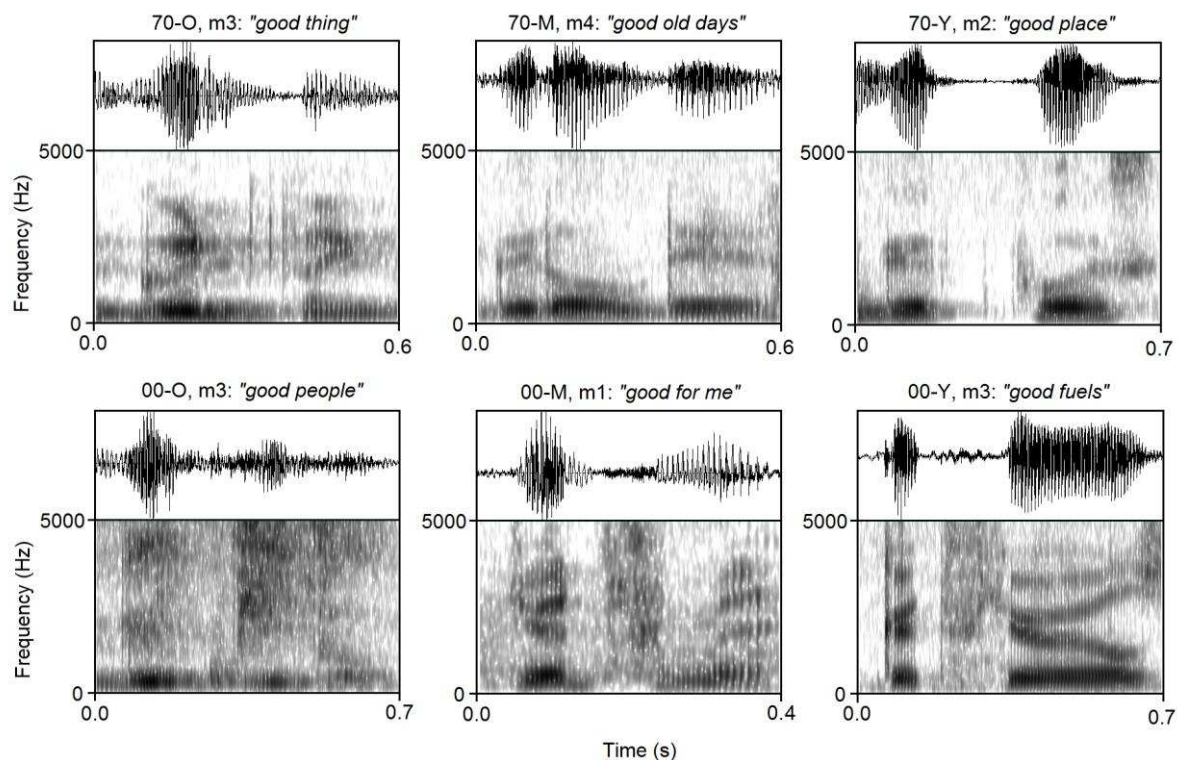


Figure 1: Waveforms and spectrograms of six samples to demonstrate differences in spectral detail across the real-time corpus of Glaswegian, 1970s recordings (upper panel) and 2000s-recordings (lower panel).

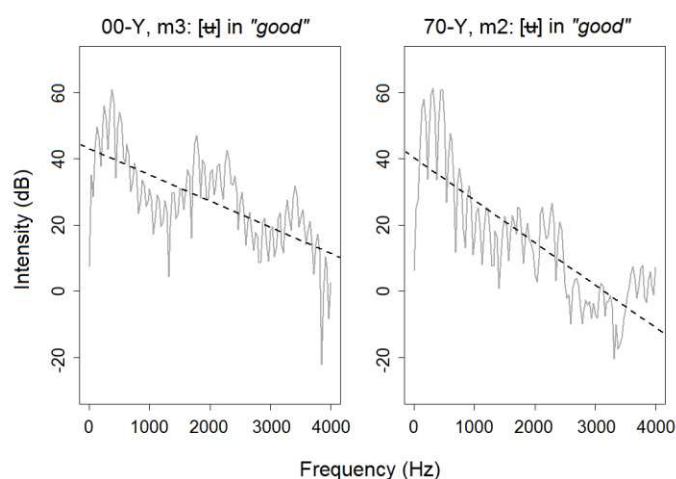


Figure 2: DFT-spectra of two [ʌ]-tokens taken from the midpoint of the vowel in 'good' (left: 70-Y, m2; right: 00-Y, m3). Dashed lines indicate spectral slope calculated as best fit by least squares regression. The spectra are based on an unsmoothed narrow-band spectrum created with a frame shift of 5 ms and a 1024 point Blackman window and then converted to a power spectrum in dB across the frequencies from 0 to 8 kHz (i.e. half the sampling rate).

To illustrate the core technical issues of interest here, Figure 3 gives an overview of spectral slope and SNR levels within the chosen set of recordings. Slope was calculated from the spectral data extracted from vowel midpoints by linear regression models in R and averaged across all vowel tokens. As indicated in Figure 2, the lower the resulting value, the steeper the negative slope, the less balanced the spectrum. To gain an insight into the levels of background noise which often do not remain constant throughout a recording, ten pauses (with the mean duration of 450 ms) were taken from various time points of each recording, mostly near the beginning, the middle and the end of each conversation (cf. Švec and Granqvist 2010). Filled pauses like those containing breathing, laughing, speech of the respective interlocutor and the like, were excluded. An average Root Mean Square amplitude (RMS, in dB) was measured for each pause and each vowel token (i.e. a RMS average was calculated across the whole duration of the respective segment). The SNR was calculated as the power ratio between the speech signal (here, vowel tokens) and the background noise (here, pauses):

$$(1) \text{ SNR} = \frac{P_{\text{signal}}}{P_{\text{noise}}} = 20 \log_{10} \left( \frac{A_{\text{signal}}}{A_{\text{noise}}} \right)$$

where P represents the average power and A the amplitude. This way of conceptualising SNR defines identical levels of noise and signal (i.e. a difference of 0 dB) as equal to 1. Accordingly, the higher the resulting SNR-value, the stronger the signal stands out against the noise. The quality of each recording was then described through its mean SNR and a mean slope value (see Figure 3).

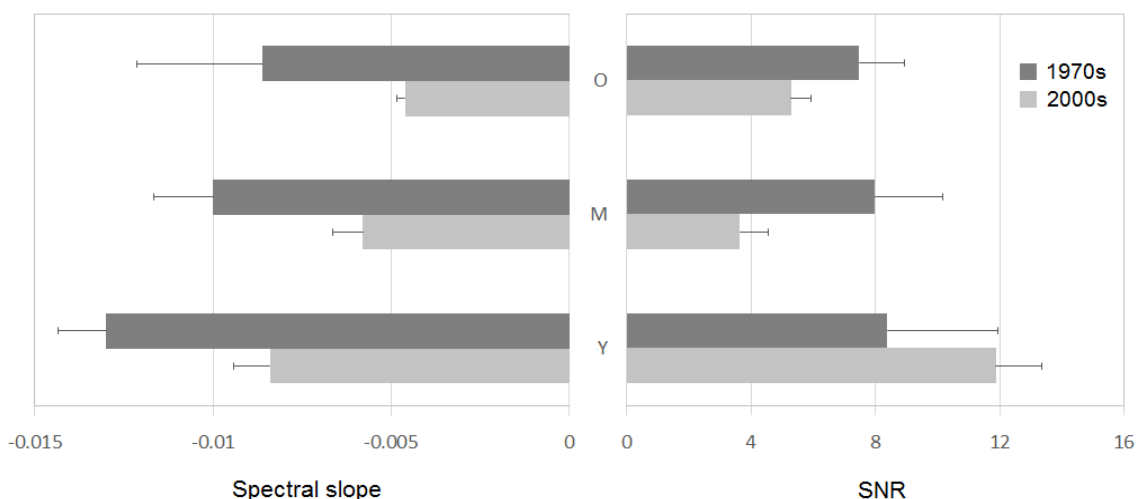


Figure 3: Mean SNR and spectral slope values and their standard deviations measured in recordings of six speaker groups (24 recordings in total, see text for more detail). The decade of recording is indicated by the two shades of grey, age group is plotted along the y-axis.

The bar graphs in Figure 3 display considerable differences in both technical aspects of the corpus recordings. By and large, the SNR was the best in the *00-Y* and the poorest in the *00-*

*M* group. All recordings from the 1970s had a more negative spectral slope in comparison to those from the 2000s. These recording-specific measurements were subjected to a series of statistical tests. The assumption of equal variances could not be accepted for a large subset of these data. Welch t-tests for two independent samples were thus performed. The observed differences in the SNR and spectral balance were compared within same-age group (real-time comparisons) and across same-decade recordings (apparent-time comparisons). The alpha-level was adjusted to account for repeated comparisons, i.e.  $p$  was set to 0.0056 ( $\approx 0.05/9$ ) in order to be interpreted as significant.

With regard to SNR, real-time comparisons did not show significant effects at the set alpha-level; though the middle-aged speaker recordings showed a trend toward significance (*70-M/00-M*:  $t_{(4.0)} = 3.7$ ,  $p = 0.021$ ), supporting the observation that recordings made with the *00-M* group (in pubs) were noisier than those made with the *70-M* group (in quiet surroundings). Apparent-time comparisons showed that SNR was significantly better in the young speakers from the 2000s in comparison to all other age groups (*00-Y/00-M*:  $t_{(5.0)} = 9.4$ ,  $p < 0.001$ ; *00-Y/00-O*:  $t_{(4.1)} = 8.1$ ,  $p = 0.0012$ ). Interestingly, there was no significant difference between recordings of the *00-M* and *00-O* groups, although their relatively poor SNR-levels result from two different sources: a strong background noise during the recording in the former vs. a distant placement of the microphone in the latter.<sup>3</sup> SNR did not differ statistically across the recordings of different age groups made in the 1970s. Neither microphone placement nor background noise levels seemed to have varied much when the older recordings took place. All of them showed a relatively high SNR, i.e. speech was relatively clear.

As far as the spectral balance is concerned, all recordings from the 1970s had a less balanced spectrum than all recordings from 2000s ( $t_{(18.9)} = 4.5$ ,  $p < 0.001$ ), which may be primarily reflective of the type of equipment used. Further real-time comparisons showed significant differences for young and middle-aged speakers (*70-Y/00-Y*:  $t_{(5.6)} = 5.0$ ,  $p = 0.003$ ; *70-M/00-M*:  $t_{(4.4)} = 4.9$ ,  $p = 0.006$ ), but the difference in the older speakers' recordings did not reach significance at the set alpha-level. These results support the above observation that an imbalanced slope seemed to be primarily an important technical specification of the recordings made by Ronald Macaulay in the 1970s (Macaulay 1977). There was only one significant apparent-time comparison, for the groups whose data were collected using different equipment (*00-Y/00-O*:  $t_{(3.3)} = 7.2$ ,  $p = 0.0039$ ). Taken together, the above findings corroborate the idea that spectral slope is related to specifications of the recording equipment (different in the 1970s vs. 2000s recordings) while SNR reflects the particulars of the recording environment (recordings made in pubs vs. quiet surroundings).

By and large, these analyses showed that from a sociolinguistic viewpoint the most relevant comparisons of the corpus (i.e. those in real-time and apparent-time as outlined

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<sup>3</sup> We also analysed the spectral properties of noise, looking for correlates of different types of noise (cf. De Decker 2016) but failed to demonstrate any significant differences in these data.

above) differed in terms of the recording quality. In the following, we will discuss potential consequences of these technical differences for acoustic analyses of vowel quality using formant analysis of the linear predictive coding (LPC).

### **3. ANALYSES**

#### ***3.1. Obtaining formant measurements***

Formants were measured at the midpoint of the vowel to reduce coarticulatory influences due to abutting segments. Formant values were extracted using standard settings of the LPC-algorithm *Burg* implemented in Praat (Anderson 1978; Boersma and Weenink 2013). By default, acoustic signals are downsampled to 10 kHz, low-pass filtered at 5 kHz with an LPC-order of 10 and a pre-emphasis of 50 Hz (i.e. starting at 50 Hz, higher frequencies are amplified by 6 dB per octave). The standard analysis window length is 25 ms with 25% window shift. Given the diverse quality recordings like those constituting the real-time corpus of Glaswegian, the low sampling rate of 10 kHz seemed advantageous for several reasons. First, it made the amount of spectral information processed by the LPC-algorithm more comparable across the two decades, thereby minimising the effects of low levels of high frequency energy in the 1970s recordings. Also, the downsampling highlights the importance to the frequency maxima below 5 kHz – a frequency range which is known to contain the most crucial information for the perception of vowel quality (Dang and Honda, 1997; Ladefoged 1962).

Our preliminary investigations showed that formants measured with these Praat default settings produced the lowest number of error rates for F1/F2 as compared to those produced by two alternative systems, EMU (Cassidy and Harrington 2001; Harrington 2010) and SFS (Huckvale 2000). The definition of error rates was roughly based on formant values measured in previous investigations of Glaswegian read speech (Stuart-Smith et al. 2013), and allowed for a generous amount of variation within 250 Hz for F1 and within 400 Hz for F2. Expected values of /i/ fell in the range of 200-450 Hz (F1) and 2000-2400 Hz (F2). F1-values outside the range of 400-650 Hz (for /u/) and 700-950 Hz (for /a/) as well as F2-values outside the range of 1300-1700 Hz (for /u/) and 1200-1600 Hz (for /a/) were considered erroneous for the purpose of the algorithm comparison. Whereas over 50% of F1/F2-values extracted from EMU and SFS were classified as erroneous in this setting, the output from Praat contained a significantly smaller amount of outliers.

#### ***3.2. Dealing with formant outliers***

Praat produced a relatively high amount of reliable measures by default but its performance could not be further improved by changing the default settings (e.g. increasing the LPC-order or altering the sampling rate). As an alternative to the chosen measurement at vowel mid-points, 'dynamic' means and medians were tested, i.e. a mean or a median value of F1 and F2 was extracted from the formant track around the central 50% of the total vowel duration. While 'dynamic' means seemed to slightly increase the

influence of segmental context on formant values, ‘dynamic’ medians led to a higher spread of the formant distributions, most considerably so for F2 ( $\pm 400$  Hz compared to the ‘static’ values) and slightly less so for F1 ( $\pm 60$  Hz compared to the ‘static’ values). Although taking formants from vowel midpoints was found to be the best method of formant extraction, the output of this ‘static’ measure still retained outlier values. Two ways of dealing with outliers were piloted on a subset of the data: manual correction and automatic outlier removal. For the manual correction of the outliers, the formant values were taken from the DFT-spectrum (512 points) and compared to the formant tracks in Praat spectrograms. Only plausible values derived from the DFT-spectrum were then included in the corrected dataset. In addition to being an extremely time-consuming procedure, this method risked introducing random variability into the sample since corrections of the same vowels by two independent experts (the first and the second author) showed disagreements within a range of  $\pm 60$  Hz for F1 and  $\pm 450$  Hz for F2. Accordingly, a more reliable procedure of data trimming by removing defined outliers was adopted instead.

The values in Hz were first Bark-transformed to map them into an auditory scale (Traunmüller 1990), then plotted and closely examined. Subsequently, F1-values above 5.5 Bark were treated as outliers for the two high vowels and those below 5.5 Bark as outliers for the open vowel. F2-values lower than 11.5 Bark for /i/ and lower than 10 Bark for /ʌ/ were excluded. Most outliers occurred in F2 of /a/, defined by the range of 9-12 Bark. Note that this approach to outlier removal is superior to a statistical trimming of F1/F2 distributions because error variances are not normally distributed so that outliers do not always fall into the upper or lower quartiles of the respective formant distribution. The overall sample size was reduced by about 15% as a consequence of the outlier removal (remaining  $n = 3296$ ).

### **3.3. Normalising raw formant data**

The Bark-scale does not serve as a means of minimising individual influences on formant measurements, and was expected to be similar to raw data in Hz (Adank, Smith and van Hout 2004). Additionally, we calculated speaker-based transformations of raw Hz values following the procedures of Lobanov (z-Hz), Nearey (N-Hz) and Watt & Fabricius (WF-Hz).

The well-established Lobanov’s and Nearey’s procedures are known to lessen the influence of speaker-specific attributes on formant values (e.g. Adank, Smith and van Hout 2004; Fabricius, Watt and Johnson 2009; Watt and Fabricius 2002). Lobanov’s approach employs z-score transformation as a means of formant normalisation (Lobanov 1971). Resulting values are the distance between each given measurement and the formant mean in numbers of standard deviations:

$$(2) F_n^{\text{Lobanov}} = \frac{F_{ni} - \bar{F}_n}{s_n}$$

where  $F_{ni}$  is an individual formant value,  $\bar{F}_n$  speaker’s mean frequency and  $s_n$  their standard deviation across all formant measurements.



In the most popular version of the Nearey's method, log-transforms of formant values are taken and the mean frequency is subtracted to derive for each measurement its distance from the mean in log-frequency units (Nearey 1978):

$$(3) F_n^{\text{Nearey}} = \log F_{ni} - \log \bar{F}_n$$

The method works at its best if (3) is applied to each formant individually (Adank, Smith and van Hout 2004). Nearey and Lobanov normalisations usually involve more than one vowel and are therefore frequently classified as 'vowel-extrinsic'. In contrast to a widespread belief, there is no imperative to sample all vowels of the system in order to achieve a vowel normalisation, but a meaningful subset is necessary. Nearey (1978: 88) suggests "*at least two points of known phonetic quality from which a speaker's formant ranges may be estimated*". The mean-based linear transformation by Nearey could work well if at least the corner vowels of the system are sampled; such procedure is likely to produce similar normalised values in comparison to a whole-system sampling approach.<sup>4</sup> In contrast, the normalised output of the Lobanov's transform very much depends on the number and the spread of the vowel categories, as its scaling unit is the standard deviation. Crucially, the same vowel categories ought to be sampled across speakers to achieve comparability in scaling which is core to a successful normalisation.

A more recently developed normalisation procedure proposed by Watt and Fabricius (2002) and subsequently refined by Fabricius, Watt and Johnson (2009), is becoming increasingly popular. The method is based on the same three principles as above, i.e. it is applied formant- and speaker-intrinsically but vowel-extrinsically. The procedure first seeks to establish geometric centres of gravity ( $S_1$  and  $S_2$ ) in the speaker's F1/F2 plane as described by three corner vowels: a close-front /i/, an open-central /a/ and a (hypothetical) close-back /u/. Each formant value is then divided by the individual normalisation constant  $S_1$  or  $S_2$ , as appropriate:

$$(4) S_n = \frac{\bar{F}_{n(i)} + \bar{F}_{n(a)} + \bar{F}_{n(u'')}}{3}; F_n^{\text{Watt\&Fabricius}} = \frac{F_{ni}}{S_n}$$

Here,  $u''$  denotes that the F1 and F2 values for /u/ are not observed but constructed. More specifically, they are set equal to the mean F1-frequency of the close front vowel /i/. Similarly, only the actual F1-mean of a speaker's realisations of the open vowel /a/ is used to calculate the  $S_1$ -constant (hence,  $a'$ ). For  $S_2$ , F2 of /a/ is interpreted as equidistant between F2 of  $i$  and F2 of  $u''$  (i.e. it equals the F1/F2-mean of /i/). In sum, the method only requires an input of the speaker's mean frequencies for F1 and F2 of /i/ and F1 of /a/ in order to provide a normalised value for any vowel.

Formant values resulting from (4) indicate how far each speaker's vowel is from their centre of gravity. In contrast to Lobanov's and Nearey's procedures, the primary motivation behind the Watt-Fabricius approach was to create a means for visual inspections of vowel spaces in speakers of different sexes, common in sociolinguistic

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<sup>4</sup> Cf.  $\text{mean}(1, 5) = \text{mean}(1,2,3,4,5)$

research. According to previous research, the method helps to reduce a considerable amount of data dispersion due to anatomical differences between speakers and has been shown, along with the Lobanov approach, to outperform Nearey on at least this criterion (Fabricius, Watt and Johnson 2009; Watt and Fabricius 2002).

### 3.4. Evaluating the performance of normalisation methods

To gain an appreciation of each method's performance in the context of technically diverse data, two measurements were obtained for each speaker's recording: (1) a measure of the overall vowel space size and (2) a measure of the dispersion within each vowel category.<sup>5</sup> Using (1), we captured the potential consequences of artificially higher F1 and/or lower F2 reported in the literature (Byrne and Foulkes 2004; De Decker and Nycz 2011; Künzel 2001; Rozborski 2007; van Son 2005). Using (2), we could estimate how well a method dealt with technically introduced dispersion of formant values (cf. Fabricius, Watt and Johnson 2009).

The size of the F1/F2-vowel space, constituted by the corner vowels /i ʌ a/, was calculated as the area of a triangle,  $A_t$ , using Heron's formula (Heath 1921, 321ff.):

$$(5) A_t = \frac{1}{4} \sqrt{(a+b-c)(a-b+c)(-a+b+c)(a+b+c)}$$

where  $a$ ,  $b$  and  $c$  are Euclidean distances between [i]/[a], [i]/[ʌ] and [a]/[ʌ], respectively, in the two-dimensional formant space. Mean values of each speaker's F1 and F2 frequencies per vowel category were used to calculate their individual  $A_t$ .

Given that vowel quality distributions are usually conceptualised as 95% confidence interval ellipses around F1/F2 means, we measured the dispersion as the area enclosed by an ellipse,  $A_e$ , using:

$$(6) A_e = \pi ab$$

where  $a$  and  $b$  are 1/2 the ellipse's major and minor axes, respectively (cf. Disner 1980). The R-package `siar` was used to calculate  $A_e$ . Figure 4 shows a schematic representation of the two measurements,  $A_t$  and  $A_e$ .

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<sup>5</sup>Previous studies sometimes utilized the squared coefficient of variation (calculated  $(SD/mean)^2$ ) as a measure of vowel dispersion (e.g. Fabricius, Watt and Johnson 2009:243). The coefficient is meant to evaluate the success of several vowel normalization methods in reducing the speaker-induced variability, and to make this evaluation independent of the scaling unit of each normalization method. A potential issue with this approach lies in the fact that the formula is inapplicable if the denominator (i.e. the mean) equals 0 which is theoretically possible in the case of Lobanov and Nearey transforms.

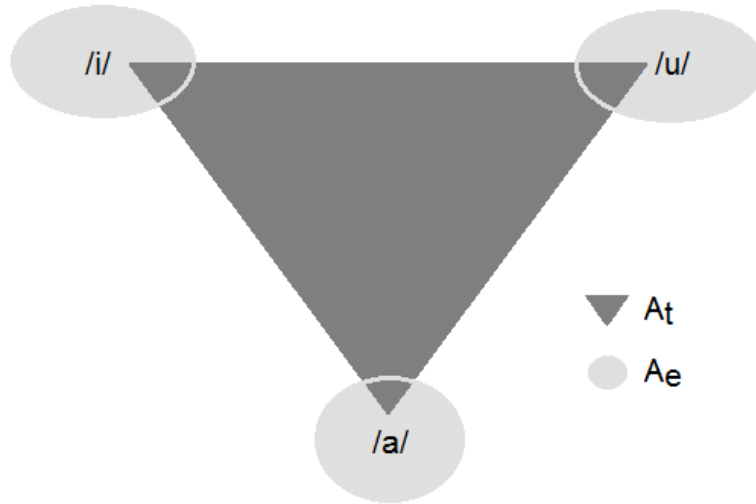


Figure 4: A schematic representation of the triangular F1/F2 vowel space as created by the corner vowels /i a u/,  $A_t$  (in dark grey) and the dispersion ellipses of 95% confidence intervals,  $A_e$  (in light grey).

The  $A_e$  metric could not be applied to /a/ of the speaker *70-Y, m01* as only 2 (out of 33 labelled) cases remained in his dataset after outlier removal. This problem resulted from erroneous tracking of F1 in many /a/-tokens of this speaker, which mostly had values as low as in speaker's /i u/ tokens. Consequently, the discussion of  $A_{e(a)}$  below will be based on the results of 23 speakers.

Subsequently, correlations between SNR and spectral slope values on the one hand and  $A_t$ ,  $A_e$  measurements of a recording on the other hand were run to uncover potential relationships between the technical quality measurements and the properties of the corresponding vowel spaces (in Bark, z-Hz, N-Hz, WF-Hz). A significant correlation would indicate that there is a linear relationship between a measurement of technical quality and a measurement of the vowel space.

### 3.5. Using an alternative normalisation by vowel distance ratios

There are further alternatives to the three normalisation procedures discussed above. Speaker normalisation is implicit in a relative measure which conceptualises ratios instead of scaled units of measurement. For investigations of sound changes affecting the fronting of /u/ in the vowel space,  $d_u$ , the logarithmic Euclidean distance ratio defines the relative positioning of the target vowel between two meaningfully chosen reference vowels as e.g. the front /i/ and the back /a/ in Southern British English (Harrington, Kleber and Reubold 2008):

$$(7) d_u = \log(E_{u/a}) - \log(E_{u/i})$$

where  $E_{u/a}$  and  $E_{u/i}$  are Euclidian distances between the respective vowels. Accordingly, mean acoustic values of each speaker's corner vowels /i a/ are treated as centroids of a multi-dimensional vowel space. The relative distance of /u/ between the two centroids is measured: when  $d_u$  is zero, the token is equidistant between the two centroids; positive  $d_u$ -

values indicate its proximity to /i/ (i.e. fronting of /u/) while negative  $d_u$  shows that /u/ is closer to /ɑ/ than to /i/ (i.e. backing of /u/). Since these ratios are calculated separately for each speaker (i.e. relative to speaker-specific centroids), then a certain degree of speaker normalisation is implicit in the calculation.

Crucially, the choice of the reference vowels needs to be considered in the context of each given variety and research question of the investigation. Southern British English back, open vowel /ɑ/ is missing in the phonology of Standard Scottish English. Another back vowel, /ɒ/, is known to be unstable and potentially undergoing fronting itself (Watt and Tillotson 2001). Since the potential sound change in question involves a lowering of Scottish /ɜ/ (which is already considerably front, see 1.3), the definition of two anchor points in the vowel space - the lowest, central /ɑ/ and the highest, front /i/ - was considered the most suitable approach for the purposes of the present investigation.

However, the core premise of the method that the reference vowels are reliable anchors may be still, at least partly, violated in our real-time corpus due to the technical issues. For example, it is possible that the frequency of front vowels with a high second formant, like our reference vowel /i/ here, is more strongly affected by the attenuation of spectral energy in the upper frequency range attested in a large part of the 1970s-recordings (Sect. 2.3): if this is the case, we might find that F2 of /i/ is lowered in those recordings while F2 of /ɜ/ or /ɑ/ remains fairly unaffected. This constellation would lead to an inherent, technically introduced bias towards a positive  $d_u$ -output (i.e. a larger proximity of /ɜ/ to /i/). If, in contrast, the issues arising from a recording-specific slope or SNR have a very similar effect on formants of both high and low vowels, the  $d_u$ -metric may offer the best approach to tackling both technical and speaker-related issues as the ratios are calculated for each recording and speaker separately. An additional argument in favour of choosing /ɑ/ and /i/ as anchors lies in the fact that the space between the centroids of the lowest and the highest vowel of the Scottish vowel system may be used to investigate a potential lowering of /ɜ/ (see 1.3).

Again, correlations between SNR, spectral slope and averaged  $d_{u(F2)}$ ,  $d_{u(F1)}$  measurements of recordings will help to shed light on whether or not these measurements co-vary;  $d_u$ -values can further be subjected to statistical analyses and serve as a dependent variable (see 3.6).

### ***3.6. Dealing with individual variation and technical issues by statistical means***

Finally, another way of dealing with various influences on formant values is to use the estimation procedure of linear mixed-effects modelling. A potential advantage of this normalisation method lies in that the speaker (reflecting individual influences) and the quality of a recording (reflecting technical issues) can be defined as random or fixed factors, allowing for their individual impacts on the dataset to be estimated. Estimates are preferable to raw means as they represent weighted means obtained in a situation when all other sources of influence defined in the model are held constant. Estimation seems to

be particularly appealing in contexts of unbalanced datasets which are common in analyses of spontaneous speech, mainly because multiple sources of influence can be accounted for in a model (cf. Hay 2011:212f). Plotting estimates instead of raw means might therefore allow substantially increased comparability of formant plots.

Linear mixed-effects models were fitted to F1 and F2 data separately using the `lme4`-library in R. Overall, there were eight dependent variables (and models) since for both formants, one model was fitted to each of the four Hz-derivatives (Bark, z-Hz, N-Hz and WF-Hz). The best model fit was established through model comparisons using `drop1()` function implemented in the R-library `lmerTest`. Each model contained a simplified structure of fixed and random effects in order to maintain some comparability with the three vowel transforms discussed above (e.g. potential effects of consonantal environment or lexical items were not considered). The quality of each recording was specified in terms of its SNR and spectral slope which were both converted to binary factors using median split.<sup>6</sup> In both cases, just over 50% of all measurements - 51% (SNR, median: 6.79) and 54% (slope, median: -0.0085) – were assigned to the “higher quality” category.

Speaker *group*, vowel category, SNR and spectral *slope* were fitted as the fixed factors, and individual *speaker* and the recording *source* as random effects. We tested for all possible, meaningful two-way interactions of the predictors.

## 4. RESULTS

### 4.1. Recording-based analyses

#### 4.1.1. Correlations between technical quality and vowel space measurements

First, correlations were run to explore the potential interdependence between SNR or spectral slope features of a recording on the one hand and its vowel space size  $A_t$  or vowel dispersion  $A_e$  on the other. We did not find significant correlations between SNR and  $A_t$  (on any of the Hz-transforms) suggesting that these phenomena are unrelated. However, there was a significant positive correlation between the spectral balance and the vowel space size showing that the more negative the spectral slope of a recording, the smaller its vowel space in Bark ( $R = 0.45$ ,  $t_{(22)} = 2.4$ ,  $p = 0.027$ ), N-Hz ( $R = 0.44$ ,  $t_{(22)} = 2.3$ ,  $p = 0.027$ ) or WF-Hz ( $R = 0.46$ ,  $t_{(22)} = 2.4$ ,  $p = 0.024$ ). This correlation was removed only if the vowel space was created via the z-Hz scale ( $R^2 = 0.1$ , n.s.).

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<sup>6</sup> The decision to create binary factors was made since only two technical specifications (one for SNR and one for slope) were available for each recording ( $N = 24$ ) that measured many tokens ( $N = 3296$ ). Hence the technical specifications could not be fitted as true covariates (which would require one technical specification per token with  $N = 3296$ ), and little advantage was seen in including two 24-level predictors into each model (given the lack of a hypothesis related to the 24 levels, and also how difficult it is to interpret the meaning of statistical significance in a multilevel factor, see Baayen 2008:114).

With respect to the amount of dispersion  $A_e$  and the recording quality, absent correlations between the  $A_e$  metric (for any of the vowels and scales) and spectral balance was suggestive of the two factors being rather unrelated. No significant correlations were obtained for SNR and  $A_e$ , either.

#### 4.1.2. Correlations between technical quality measurements and distance metric $d_u$

Next, we ran correlations between SNR or spectral slope values and the speaker-specific averages of the distance metrics  $d_{u(F2)}$  and  $d_{u(F1)}$ , indicative of the degree of /ʌ/-fronting and -lowering, respectively. A marginal effect was found for SNR and  $d_{u(F1)}$ , showing a negative correlation: the lower SNR-levels (i.e. the noisier the recording), the higher the  $d_u$  values, i.e. the closer /ʌ/ tends to move to /i/ in the created vowel space ( $R = -0.37$ ,  $t_{(22)} = -1.9$ ,  $p = 0.076$ ). There were no further effects.

### 4.2. Token-based analyses

#### 4.2.1. Linear mixed-effects statistic for F1

Linear mixed-effects models (see 3.6) were fitted to the measurements of F1 (in Bark, z-Hz, n-Hz, WF-Hz). The outputs of the best fit model for each of the Hz-transforms are shown in Table 1. Two core interactions were significant regardless of the scale, one potentially indicative of a technical issue impacting on F1 (*vowel\*slope*) and one related to the core sociolinguistic interest behind formant analyses – a potential sound change (*vowel\*group*).

Table 1. Output of linear mixed-effects models for F1-values through four different Hz-derivatives (Bark, z-Hz, n-Hz, WF-Hz).

Hz-derivate	Factor/interaction	$\chi^2$	df	p
Bark	<i>Vowel*SNR</i>	13.9	2	<0.0001
	<i>Vowel*slope</i>	144.9	2	<0.0001
	<i>Vowel*group</i>	532.1	10	<0.0001
WF-Hz	<i>Vowel*slope</i>	107.75	2	<0.0001
	<i>Vowel*group</i>	699.45	10	<0.0001
z-Hz	<i>Vowel*slope</i>	37.9	2	<0.0001
	<i>Vowel*group</i>	74.5	10	<0.0001
n-Hz	<i>Vowel*slope</i>	93.6	2	<0.0001
	<i>Vowel*group</i>	626.7	10	<0.0001

With respect to the technical influences, vowel quality interacted with spectral slope. Post-hoc t-tests of the best-fit models compared the impact of the slope imbalance on the F1-values measured in /a/, /i/ and /ʌ/ separately and showed noteworthy discrepancies between the four scales. On the Bark scale, an effect of the slope imbalance was observed exclusively in the open vowel /a/ whose F1 was raised by 0.57 Bark when the slope was more negative ( $t_{(10)} = 3.4$ ,  $p < 0.01$ ); a significant effect was absent in the two high vowels /i/ and /ʌ/. The spectral balance effect on F1 of /a/ was merely trending toward significance after the frequency transformation using the Nearey formula ( $t_{(17)} = 1.8$ ,  $p = 0.088$ ) while F1 of

/i/ and /ɨ/ remained unaffected by the slope changes. Interestingly, this effect on /a/ disappeared completely after the Lobanov-transform (still no effects were observed for /i/ or /ɨ/). In contrast, the Watt-Fabricius transform uncovered a strong effect of /a/ having a higher F1 when the spectral slope was less balanced (0.06 higher,  $t_{(19)} = 4.2$ ,  $p < 0.001$ ). Additionally, the slope also had an impact on F1 measured for /i/ (0.07 units lower,  $t_{(19)} = 3.0$ ,  $p < 0.01$ ) and /ɨ/ (0.06 units lower,  $t_{(19)} = 2.6$ ,  $p < 0.05$ ). This result for high vowels may point to the central role of /a/ as a reference vowel in this transformation (see (4)). We will return to the discussion of these findings in Sect. 5.3 below.

In addition to the effect of the spectral slope, models fitted to Bark-scaled F1-values also showed a significant interaction of the vowel quality and SNR. Low SNR-levels (meaning less clear recordings) raised F1 of all vowels, though the effect was the strongest for /a/ (0.57 Bark,  $t_{(18)} = 4.0$ ,  $p < 0.001$ ), slightly less distinct for /i/ (0.44 Bark,  $t_{(18)} = 42.9$ ,  $p < 0.01$ ) and even weaker for /ɨ/ (0.35 Bark,  $t_{(18)} = 2.5$ ,  $p < 0.05$ ). The interaction was not significant for any other Hz-transforms.

Second, vowel quality also interacted with speaker group for all Hz-transforms (see Table 1). If the age-related differences were completely accounted for by the normalisation procedures (see 3.3), this finding may be suggestive of a sound change (we will address this question in 4.2.4-5).

#### 4.2.2. Linear mixed-effects statistic for F2

Subsequently, another series of linear mixed-effects models was run for the F2-values (in Bark, z-Hz, n-Hz, WF-Hz). The outputs of the models with the best fit are outlined in Table 2. All models showed an effect potentially related to sound change (*vowel\*group*) which will be addressed in 4.2.4-5. Models for Bark, Watt-Fabricius and Nearey scales further showed two interactions indicative of an influence of the technical issues in the recordings (i.e. *vowel\*SNR* and *vowel\*slope*). In contrast, Lobanov-transformed F2-values (z-Hz) did not display the effect of the spectral balance, and only produced a comparably weak effect of the different SNR-levels. Subsequent t-tests, however, failed to produce a significant effect among the relevant contrasts between more vs. less noisy recordings for /i ɨ/ or /a/. Similarly, these contrasts were not significant in the t-tests run for n-Hz scale, meaning that both Nearey and Lobanov transforms eliminate the impact of recording quality on F2-measurements.<sup>7</sup>

On the Bark scale, t-tests produced evidence that F2 of /i/ was 0.4 Bark lower in noisy recordings ( $t_{(20)} = 2.5$ ,  $p < 0.05$ ) and 0.5 Bark lower in recordings with a less balanced spectral slope ( $t_{(18)} = 2.8$ ,  $p < 0.05$ ), but no evidence for such effects in /a/ or /ɨ/. WF-transform showed the most significant effects. Here, /a/ and /ɨ/ (but not /i/) were both affected by the two technical issues. More specifically, F2 of /a/ measured 0.09 WF-Hz

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<sup>7</sup> The significant interactions *vowel\*SNR* or *vowel\*slope* listed in Table 2 are related to other (in our case meaningless, yet routinely calculated by the linear mixed effects procedure) contrasts between the factor levels, e.g. F2 of a noisy /a/ vs. F2 of a quiet /i/.



higher values in poorer SNR ( $t_{(20)} = 3.1$ ,  $p < 0.01$ ) as well as in less balanced spectral slopes ( $t_{(10)} = 2.8$ ,  $p < 0.05$ ). For /ʌ/, F2 was 0.08 WF-Hz higher in noisy recordings ( $t_{(12)} = 2.4$ ,  $p < 0.05$ ) and 0.06 WF-Hz higher in recordings with an unbalanced slope ( $t_{(21)} = 1.9$ ,  $p = 0.066$ ). These findings are very likely to be related to the way the normalisation constant  $S_2$  is calculated in (4), and will be discussed in Sect. 5.3 below.

Table 2. Output of linear mixed-effects models for F2-values measured on Bark, z-Hz, n-Hz, WF-Hz scales.

Hz-derivate	Factor/interaction	$\chi^2$	df	p
Bark	<i>Vowel*SNR</i>	52.1	2	<0.0001
	<i>Vowel*slope</i>	57.7	2	<0.0001
	<i>Vowel*group</i>	157.2	10	<0.0001
WF-Hz	<i>Vowel*SNR</i>	45.3	2	<0.0001
	<i>Vowel*slope</i>	48.9	2	<0.0001
	<i>Vowel*group</i>	131.8	10	<0.0001
z-Hz	<i>Vowel*SNR</i>	8.4	2	<0.05
	<i>Vowel*group</i>	72.9	10	<0.0001
n-Hz	<i>Vowel*SNR</i>	52.5	2	<0.0001
	<i>Vowel*slope</i>	59.7	2	<0.0001
	<i>Vowel*group</i>	163.9	10	<0.0001

#### 4.2.3. Linear mixed-effects statistic for $d_u$ -metrics

As the next step, a mixed-effects model was fitted to the  $d_{u(F1)}$  and  $d_{u(F2)}$  metrics discussed in 3.5, which measures the location of /ʌ/ in the F1 or F2-space between the corner vowels /i/ and /a/, potentially indicative of lowering and/or fronting of /ʌ/. The new models retained the same structure of random effects as the token-based models discussed above (Sect. 4.2.1 and 4.2.2). The predictors were *group*, *SNR* and *slope*; only one interaction (*SNR\*slope*) was tested for.  $d_{u(F1)}$  and  $d_{u(F2)}$  served as the dependent variables in two separate calculations.

The best-fit models for both  $d_{u(F1)}$  and  $d_{u(F2)}$  did not contain any effect for either factor related to the technical issues under investigation. The factor *group* showed significance for  $d_{u(F1)}$  ( $\chi^2(5) = 11.4$  and  $p = 0.044$ ), but only a trend for  $d_{u(F2)}$  ( $\chi^2(5) = 10.3$  and  $p = 0.068$ ).

#### 4.2.4. Visual representations on F1/F2-plane and inference of change

Since exploratory formant plots are a common sociophonetic tool to discuss and ascertain sound variation and to derive change (e.g. Labov, 1994; Labov, Ash and Boberg, 2006),<sup>8</sup> this section is dedicated to exploring how the significant technical effects shown above might interfere with the interpretation of the visual data, and focuses on two scales: Bark and z-Hz, given that Bark-scaled values contain the information related to both the speaker physiology and the technical set-up of the recording, while Lobanov-transformed z-Hz

<sup>8</sup> Also the NORM suite facilitates this well: <http://lingtools.uoregon.edu/norm/>



values seem to retain the least influence from either factor (cf. also Adank, Smits and van Hout, 2004).

Formant plots in Figure 5 and 6 display F1/F2 values with the 75%-dispersion and the centroids of /i ʌ a/. Additionally, the graphs compare the sizes (and shapes) of the vowel spaces derived through the interpolation between the centroids of /i ʌ a/. The centroids were based either on raw-data averages (black lines) or on estimates of the best-fit linear mixed-effects models (grey lines, see 3.6). The sizes of the resulting vowel spaces were measured as an area of a triangle,  $A_t$  (see 3.4), and given for comparison.

Patterns in Figures 5 and 6 suggest two core observations. First, the triangular vowel spaces appear more unevenly sized across the six speaker groups when plotted on the Bark scale than on the z-Hz scale.  $A_t$  value of the 70-Y group was partly less than half the value of any other speaker group. In contrast, z-Hz scale created a more balanced representation of the six vowel spaces in this sample. Even though the triangle between /i/, /ʌ/ and /a/ remained the smallest in the 70-Y group under this transform, the magnitude of the differences between the groups diminished (reflected in more comparable  $A_t$  measurements across the sample). Second, the issue of the Bark-scaled vowel spaces being of highly varied sizes across the sample could not be resolved using statistical means of estimation implemented in linear mixed-effects modelling. Moreover, this method created a number of F1-values substantially diverging from raw-data means, most notably for the open vowel /a/ (i.e. the vowel whose F1 was particularly strongly affected by both technical issues under investigation, see 4.2.1). Such substantial discrepancies between the means and the estimates of formant values were absent after the Lobanov transform, reinforcing the idea that technical issues can be effectively dealt with by applying a vowel normalisation.

The relative size of the formant ellipses in Figures 5 and 6 seems to suggest an increase in F1/F2-dispersion after the Lobanov transformation. However, this visual observation is deceiving and results from differences in the units of scaling. In fact, the ellipse sizes measured as  $A_e$  (see (6) in 3.4) were slightly but significantly larger on the Bark scale than on the z-Hz scale, for /i/ ( $t_{(45.9)} = 2.9$ ,  $p=0.0065$ ), /ʌ/ ( $t_{(45.3)} = 3.2$ ,  $p=0.0024$ ) as well as /a/ ( $t_{(43.6)} = 2.5$ ,  $p=0.018$ ) while  $A_t$  values (see (5) in 3.4) did not differ significantly between the two scales.

With respect to the inference of change, statistical results (see 4.2.1-2) were a little inconclusive. Surprisingly, none of the planned group comparisons showed significance for F1/F2 measured on the z-Hz scale. On the Bark scale, some apparent-time comparisons were significant for F1 and F2; significant real-time comparisons were observed for F2 only.

As far as F1 was concerned, only 70-Y group entered significant apparent-time comparisons. Compared to the middle-aged speakers, the young speakers recorded in 1970s had a 0.5-0.8 Bark higher F1 in high vowels (70-Y/70-M-comparisons for /i/:  $t_{(18)} =$

2.7,  $p < 0.05$ ; 70-Y/70-M-comparison for /ʌ/:  $t_{(18)} = 3.4$ ,  $p < 0.01$ ) and a 0.4 Bark lower F1 in low vowels (70-Y/70-M-comparison:  $t_{(18)} = 2.6$ ,  $p < 0.05$ ).

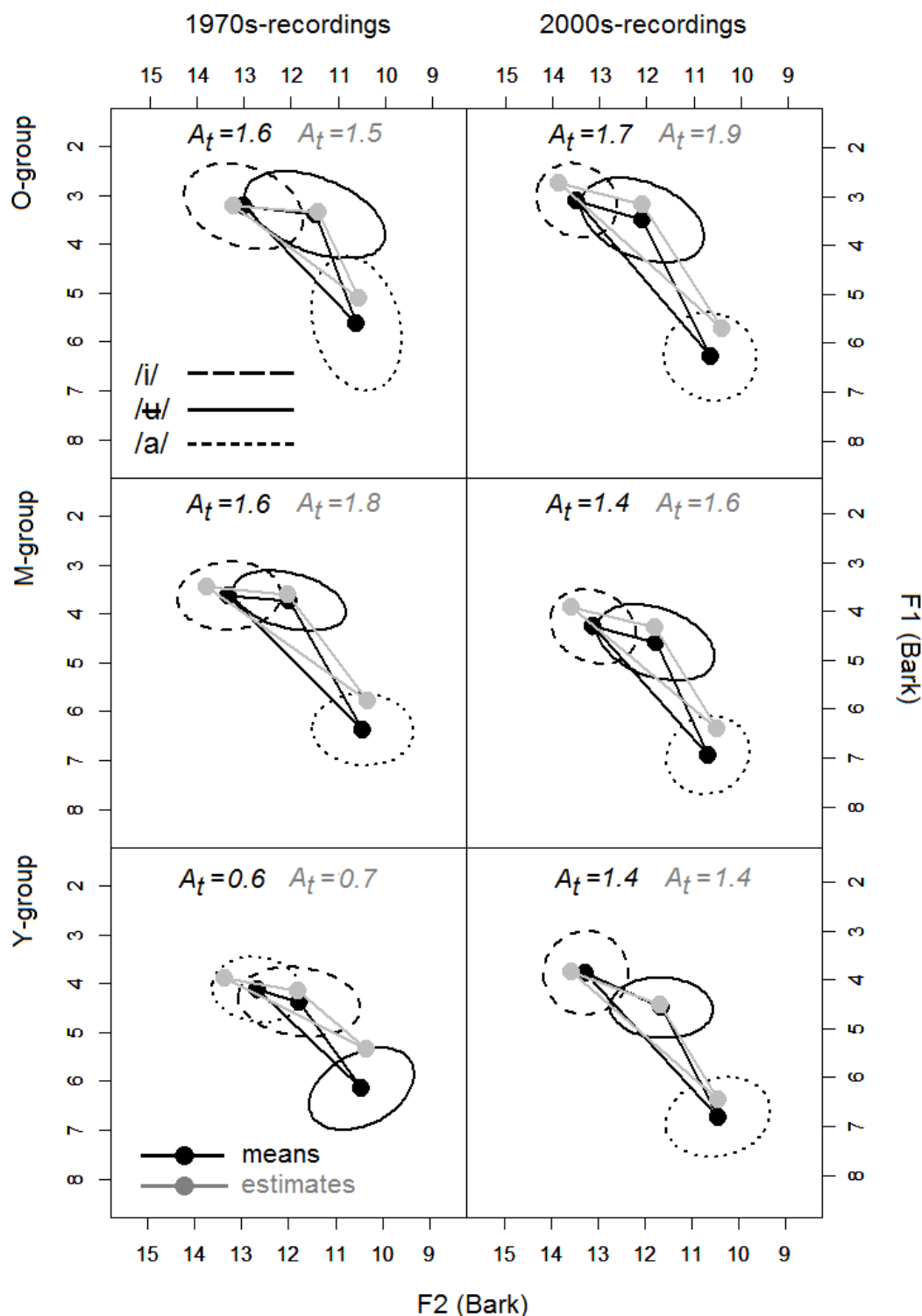


Figure 5: Means (black lines) and estimates (grey lines) of F1/F2 values of /i a ʌ/ measured in Bark. 75% confidence interval ellipses show the dispersion of individual values measured for /ʌ/ (solid lines), /i/ (dashed lines) and /a/ (dotted lines). The three age groups are plotted from top (old speakers) to bottom (young speakers). Recordings from 1970s are shown on the left, the ones from 2000s on the right. See text for further detail.

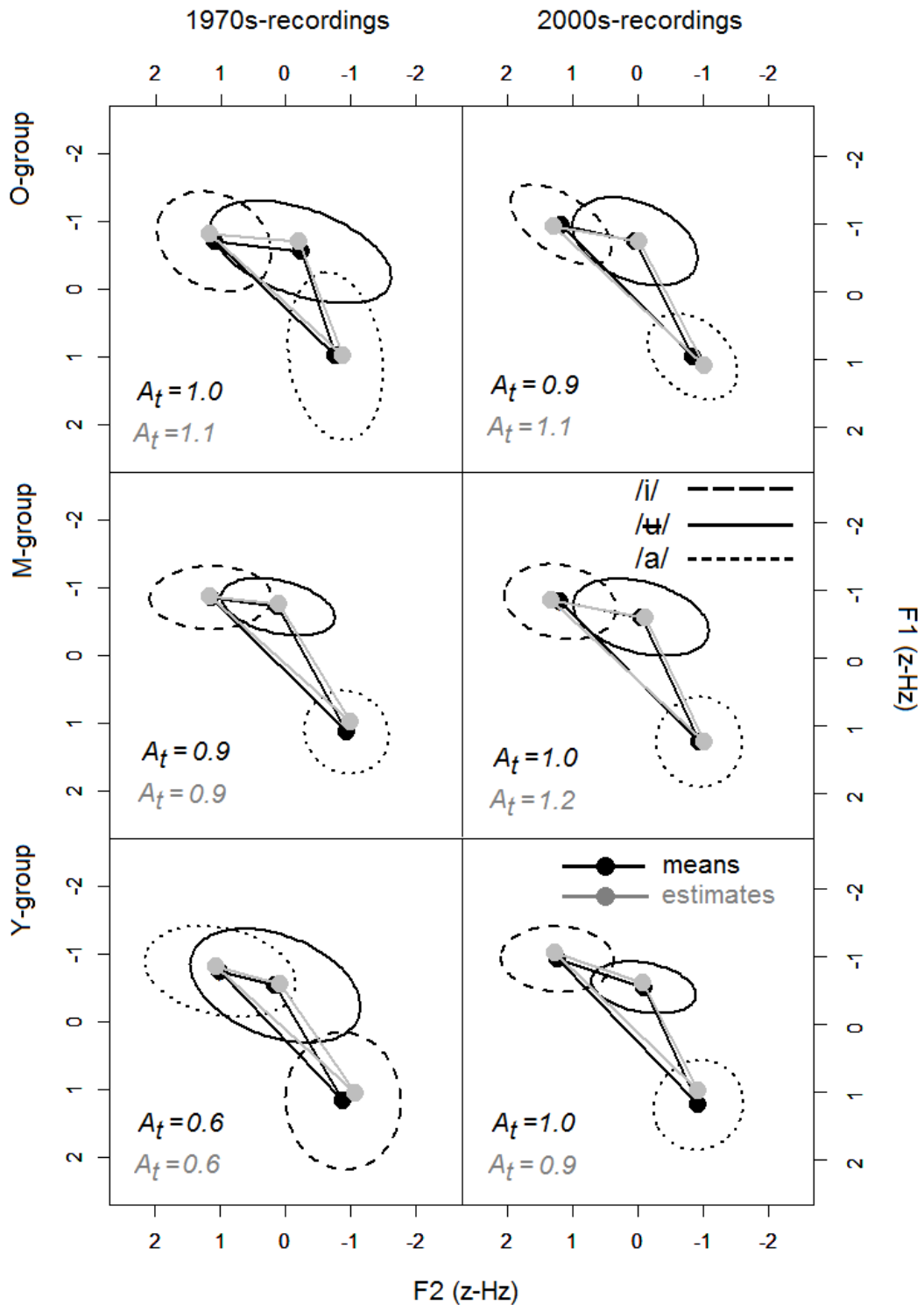


Figure 6: Means (black lines) and estimates (grey lines) of F1/F2 values of /i a ɘ/ measured in Lobanov-transformed z-Hz. 75% confidence interval ellipses show the dispersion of individual values measured for /ɘ/ (solid lines), /i/ (dashed lines) and /a/ (dotted lines). The three age groups are plotted from top (old speakers) to bottom (young speakers). Recordings from 1970s are shown on the left, the ones from 2000s on the left. See text for further detail.

Regarding F2 measurements, only high vowels of the 70-O group showed significant comparisons in real-time (70-O/00-O-comparison for /i/:  $t_{(18)} = 2.9$ ,  $p = 0.01$ ; 70-O/70-M-comparison for /ʉ/:  $t_{(19)} = 2.9$ ,  $p = 0.01$ ) and in apparent-time (70-O/70-M-comparison for /i/:  $t_{(19)} = 3.2$ ,  $p = 0.0047$ ; 70-O/70-M-comparison for /ʉ/:  $t_{(19)} = 3.6$ ,  $p = 0.0019$ ), with 70-O group measuring 0.6-0.7 Bark lower F2 of both high vowels.

While the results for F1 are rather suggestive of technically influenced patterns, the results for F2 could be cautiously interpreted as indicative of change, signifying a /ʉ/-fronting that took place in Glasgow between 1890s and 1920s (70-O vs. 70-M/00-O speakers). However, not only /ʉ/ but also /i/ shows a similar rise of F2-frequency in more recent speaker groups. Fronting of both /ʉ/ and /i/ cannot be expected under the sound change view, and prompts a question about technical issues or other recording-related factors influencing the results.

#### 4.2.5. Inference of change based on $d_u$ -measure

As seen above (4.1.2; 4.2.3), the  $d_u$ -measure was largely unaffected by the technical issues under investigation.<sup>9</sup> In contrast, the factor *group* was relevant for explaining the variation in these F1/F2 data. Figure 7 displays the group results. Despite a visible, continuous (apparent- and real-time) tendency for F1 of /ʉ/ to shift away from /i/ (i.e. to lower), planned comparisons yielded no significant effects for  $d_{u(F1)}$  across the sample. In contrast,  $d_{u(F2)}$  showed two real-time effects involving the young (00-Y/70-Y:  $t_{(17.3)} = 2.6$ ,  $p = 0.017$ ) and the old (00-O/70-O:  $t_{(15.1)} = 2.1$ ,  $p = 0.05$ ) speakers (the comparison for middle-age speakers was merely trending towards significance with  $t_{(21.7)} = 1.9$ ,  $p = 0.072$ ).

In terms of an inferred change, the effects involving F2 point to two different directions: on the one hand, an early /ʉ/-fronting dating as far back as 1890s and 1920s (70-O vs. 00-O groups, cf. also 4.2.4); on the other hand, a more recent /ʉ/-backing – i.e. a reversal of the previous change – which may have taken place between 1965 and 1985.

The patterns for F1 point subtly, yet somewhat consistently, in the direction of a potential /ʉ/-lowering over time, and cannot be explained by stylistic factors which in fact predict a completely opposite pattern of more retracted vowels in sociolinguistic and oral history interviews (all 70s and 00-O groups) than in the spontaneous conversations (00-Y and 00-M groups, see 2.1). The fact that neither apparent-, nor real-time comparisons produced a significant effect may be due to a relatively small sample of this study and a relatively small magnitude of the change (see Stuart-Smith et al 2016). Consequently, an additional model was fit to the  $d_{u(F1)}$  data, replacing the predictor *group* by two factors, *year* of recording (00 vs. 70) and *age* group of the speakers (O, M, Y). The best-fit model included both factors (*age*:  $\chi^2(2) = 7.4$ ,  $p = 0.025$  and *year*:  $\chi^2(1) = 4.8$  and  $p = 0.029$ ). In contrast to the speakers recorded in the 2000s, speakers from the 1970s had their  $d_{u(F1)}$  values 0.4 units closer to /i/ than to /a/, suggesting some vowel lowering in real-time. F1 of young speakers' /ʉ/ was 0.5 units closer to /a/ in comparison to the middle-aged ( $t_{(21.5)} = 2.2$ ,  $p = 0.035$ ) and to

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<sup>9</sup> The  $d_u$  for this investigation was calculated with Bark-scaled values (see 3.5). Noteworthy is a very strong correlation between  $d_u$  resulting from the different scales ( $R^2=0.99$ ,  $p<0.0001$ ).

old speakers ( $t_{(19.2)} = 2.4$ ,  $p = 0.028$ ), while the latter two age groups did not differ significantly from each other. These age group results are difficult to reconcile with the idea of age grading which would rather predict middle-aged speakers to deviate significantly from the remaining speaker groups (Wagner 2012). We will return to this discussion in 5.5 below.

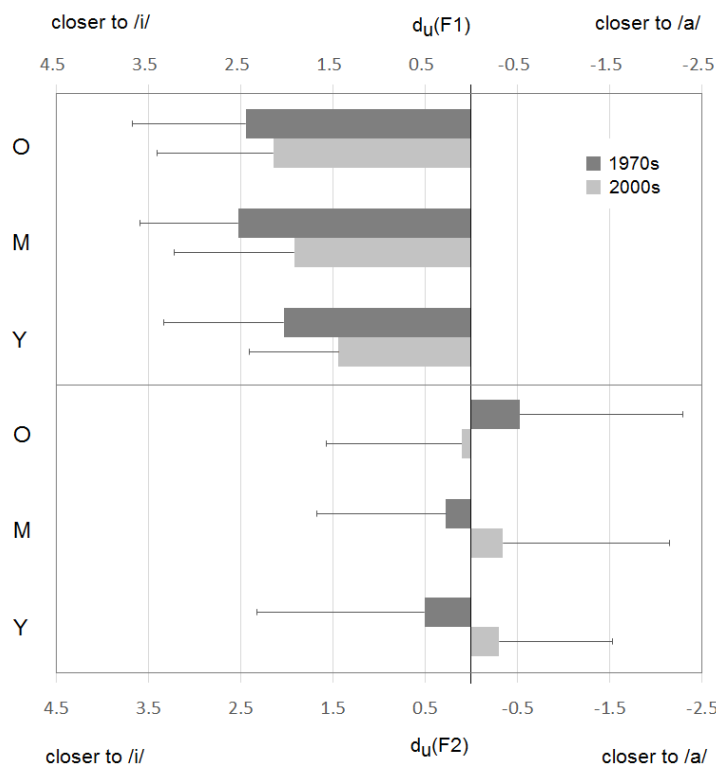


Figure 7: Group means and standard deviations of the distance metric  $d_u$  for F1 (top panel) and F2 (bottom panel).

## 5. DISCUSSION

### 5.1. Does technical quality influence acoustic measurements of vowel quality?

The primary goal of the study was to ascertain whether or not differences in recording quality – in our case, varied levels of noise, measured as SNR, and variable amounts of spectral energy available at lower and higher frequencies, measured as spectral tilt – would affect F1/F2 formant measurements in a sociolinguistic real-time corpus comprising of spontaneous speech. In this section, we will concentrate on the results obtained for the “raw” F1/F2 measurements on the Bark scale which we took as the reference point of the comparison with the normalised scales and the vowel ratio metric  $d_u$  (to be discussed below in 5.3).

To address the primary question of this study, we first confirmed that there were indeed some statistically significant technical differences among the recordings of the analysed corpus. The two technical specifications, SNR and spectral slope, showed significant differences between recording series relevant for the sociolinguistic comparisons in both real-time (here, 1970s vs. 2000s) and apparent-time (here, old vs. middle-aged vs. young

speaker recordings; see 2.3). Subsequently, we conducted recording-based, correlational analyses which unveiled the tendency for those recordings with a more negative average spectral slope (i.e. recordings which had a reduced amount of energy above 4 kHz) to produce a smaller vowel space (see 4.1.1): the more negative the slope, the smaller the resulting vowel space. Interestingly, the SNR-levels did not co-vary with either the vowel space size or the vowel dispersion – the two parameters that were measured as indicators of technically introduced artefacts on the resulting F1/F2 vowel spaces (see 3.4).

Subsequently, we looked beyond a pure co-variation of the technical specifications and vowel space measurements, and tested for causal influences of the technical quality of a recording on F1/F2. “Better” recording quality was defined by higher SNR-levels and a less tilted average spectral slope.

A significant effect of increased noise (i.e. in recordings with SNR below 6.8) in interaction with the vowel quality was observed on both F1 and F2. Noisy recordings raised F1 of all vowels, with the effect being strongest for the open vowel (amounting to an increase of approximately 0.6 Bark) and slightly weaker for the two high vowels (amounting to an increase of approximately 0.4 Bark). In contrast, only F2 of /i/ (but not F2 of /a/ or /ʌ/) was affected, and lowered by 0.4 Bark in noisier recordings. Interestingly, related forensic investigations into the impact of phone transmissions brought to light quite comparable patterns of F1-raising in all vowels across the board and a differential F2-lowering in dependence upon vowel quality (Byrne and Foulkes 2004; Künzel 2001), although in contrast to our results, the magnitude of the raising effect was usually found larger in low-F1 vowels like /i/ and /ʌ/ than in the in high-F1 vowel/a/ (whose frequency mostly lies within the transmitted range).

Spectral imbalance (with negative slopes tilted beyond -0.009) again influenced both F1 and F2 values, and interacted with the vowel quality. But for F1, we found an effect exclusively on the open vowel /a/ which measured 0.6 Bark higher values in recordings with less balanced spectra. For F2, only /i/ showed an influence, and 0.5 Bark lower formant values in recordings of poorer quality. By and large, these results are essentially in line with some of the effects reported in the previous literature (see 1.2), particularly with respect to complex interactions of technically determined factors with vowel quality (Hansen and Pharao 2006; in progress; see 5.2).

The magnitudes of the technically introduced effects are somewhat difficult to compare across the studies, given the variability in the preferred method of reporting the results – as percentage of the original frequency (Byrne and Foulkes 2004; Künzel 2001), semitones (Hansen and Pharao 2006; in progress) or raw Hz values (De Decker and Nycz 2011) – but we note that in our study, we find rather comparable magnitudes for both changes in F1 and F2, unlike previous research that showed larger deviations for F1 than for F2 which seems less affected, if at all (Byrne and Foulkes 2004; De Decker and Nycz 2011; Hansen and Pharao 2006; in progress; Künzel 2001). Such discrepancies may have various

explanations, and include the calculations (mean values based on raw data in previous studies vs. estimates from mixed-effects models in our case when random effects due to the speaker and token are accounted for, see 3.6) as well as type of the speech data (recordings of read speech in previous studies vs. spontaneous speech in our case). Moreover, our study attempted to disentangle the effects of the two influencing properties of the spectrum, SNR and spectral balance, while previous research seems likely to be dealing with both spectral features simultaneously (see 1.2). When these technical effects accumulate, their magnitude increases (see 5.2 below).

We note that the interaction of SNR and spectral slope was found significant neither for F1 nor for F2; SNR and spectral slope balance thus seem to be independent technical issues with their independent (if present) effects on the formant tracking. This finding has implications for the best-practice approaches to dealing with technically diverse recordings which we will discuss in 5.4. In sum, our analyses revealed that both noise in the recording and its spectral imbalance influence the traditional LPC-based F1/F2 formant measurements, and should therefore be not ignored in sociophonetic real-time studies that involve formant measurements as indicators of sociolinguistic variations and long-term changes in vowel quality.

### ***5.2. Are different vowel qualities affected in similar ways?***

If technical issues affected all vowel qualities in similar ways, we could have easily estimated the direction of the influence and subsequently developed a unified way of dealing with such an influence across a diverse vowel set. However, our results suggest that such a simplistic approach to dealing with technical issues of diverse recordings will remain impossible, given the persistent interaction of vowel quality with each technical issue investigated in the present study. Künzel (2001:93) arrived at a similar conclusion, faced with the vowel- and speaker-specific variability in his data.

As expected (see 1.2), corner vowels of the system were affected in particular. In this study, the high, most front vowel /i/ and the most open vowel /a/ were highly susceptible to a strong influence from both noise and spectral imbalance. More specifically, the highest F1 (in /a/) and the highest F2 (in /i/) seemed to be targeted: while F1 was strongly raised, F2 was lowered by both noise and spectral balance issues. Given the independency of both technical effects and yet the same direction of their influence, their impact on formant values accumulates (instead of e.g. cancelling each other out), raising or lowering the value by a substantial amount of up to over 1 Bark.

Although we found that corner vowels were most affected by the technical issues, we recommend real-time studies of central or mid-high peripheral vowels also run technical quality checks before attempting meaningful vowel analyses – this seems crucial since noise affected F1 of all vowels of this study, even if /ʌ/ was affected to a lesser degree. Apart from that, the advice to only investigate non-corner vowels in real-time studies of variation and change seems neither appealing nor viable.

### ***5.3. Can technical interference be effectively dealt with post-hoc?***

As shown in 1.2, the technical set-up during a recording is bound to have an impact on the resulting quality. The details and magnitudes of such technical effects are somewhat too diverse to generalise, e.g. we know that F1 is often raised in technically compromised recordings but the magnitudes of the frequency increase vary substantially across studies and (at least partly) depends on the point of comparison, i.e. on the technical specification of the recording that is considered free of such interferences. The best a researcher can do in order to achieve a high level of comparability in the sense discussed in Gregersen and Barner-Ramussen (2011) is to keep the recording equipment and surroundings exactly the same across all sessions. It may be particularly helpful to take a photo of the recording set-up if a session takes place outside of a recording studio (Christoph Draxler, personal communication). Unfortunately, researchers often neither have control over the technical set-up, nor have access to a detailed, photographic representation of the field situation during the collection of data relevant for the compilation of real-time corpora. Moreover, recording equipment is constantly evolving and being upgraded, making the exact replacement of an old, defective gadget often impossible, if the time depth between recording session is 10 or more years. Therefore, a post-hoc way of dealing with any technical influence is, and will remain, crucial to any real-time studies into sound variation and change.

In Sect. 3, we discussed an array of different, theoretically plausible approaches to dealing with technical issues, which we aligned with presently well-understood and widely-applied methods of neutralising the speaker-specific influences on F1/F2: these included three frequently applied formant normalisation procedures – Nearey (1978), Lobanov (1971), Watt-Fabricsius (Watt and Fabricsius 2002) – plus a distance ratio metric  $d_u$  (Harrington, Kleber and Reubold 2008) and statistical means (cf. Hay 2011). We will comment on how each of the above methods fared in comparison to the Bark-scaled F1/F2 data discussed above. The point of comparison in our case is the part of our corpus which has better technical specifications in terms of SNR and spectral slope balance (see 2.3).

First of all, the co-variance of the spectral slope and the vowel space size were observable for two out of the three normalised scales, namely Nearey and Watt-Fabricsius. After these transformations, the resulting vowel space remained smaller in recordings with a more negative spectral slope. In contrast, the Lobanov-transformation removed this correlation. While the SNR-levels did not seem to co-vary with either the vowel space size or the vowel dispersion measured on any of the Hz-transforms, the  $d_{u(F1)}$  values showed a potential to be affected by SNR, and a tendency to increase in noisier recordings (see 4.1.2), i.e. /ʌ/ would appear less lowered in noisier recordings. This correlation is likely attributable to the differential impact of low SNR on F1, raising that of /a/ more substantially than that of /i ʌ/, and may become more substantial in a larger corpus. In the case of our relatively small database consisting of just 24 recordings in total, however, this linear relationship was rather weak, and the correlation did not approach significance. Similarly though, the effect of SNR was absent in the linear mixed-effects models fitted to  $d_{u(F1)}$ . Overall, the distance



metric  $d_u$  seemed to provide a good way of removing the impact of technical issues from the vowel quality data since none of the technical specifications showed a significant effect on the tested  $d_u$ -values (apart from the marginally significant correlation above).

Overall, the impact of noise on F1 of /i a ʌ/ was successfully removed by all normalisation procedures, including  $d_{u(F1)}$ . The impact of the spectral imbalance on F1 of /a/ was successfully dealt with by both Nearey and Lobanov normalisations (though it is noteworthy that Nearey-scaled data retained a marginal effect). In contrast, this effect was rather boosted by the Watt-Fabricius normalisation; the transformation also led to a significant impact of spectral imbalance on both high vowels, which was absent in Bark-scaled data. These findings are very likely driven by the central role of /a/ in the calculation of the  $S_1$ -constant in the Watt-Fabricius normalisation (see (4)). If, as in our data, F1 of /a/ is the only vowel to be strongly affected by the technical issues (but not /i/, the other anchor vowel of the  $S_1$ -constant), there will be carry-over effects to the normalised F1 of any other vowels. The absence of the impact of SNR on F1 under this transformation might be due to the fact that F1 of both reference vowels, /a/ and /i/ was similarly affected by noise, raising the frequency of the formant in comparable ways: 0.6 Bark for /a/ and 0.4 Bark for /i/; the difference of 0.2 Bark between the reference vowels seemed less likely to have as strong an impact as the difference of 0.6 Bark above.

Similarly, the impact of noise and spectral balance on F2 of /i/ was successfully dealt with by two normalisation methods, Nearey and Lobanov, while the Watt-Fabricius transformation produced more significant effects than the Bark scale did. First of all, the technical issues affected /a/ and /ʌ/ (but not /i/ as in the Bark-scaled data). Instead of /i/ showing a lowering of F2 by 0.4-0.6 Bark under the influence of the technical issues, /a/ and /ʌ/ had higher F2-values after the Watt-Fabricius transformation. These patterns clearly differ from the ones ascertained for F1, and can be explained by the way the calculation of the  $S_2$ -constant works in (4): in contrast to  $S_1$ ,  $S_2$  relies exclusively on F1 and F2 of /i/. There is again a carry-over effect on F2 of other vowels if F2 of /i/ is technically affected, but it does not surface in the normalised values of /i/ itself (given that it serves as an anchor).

And finally, the success of minimising technical interference purely by the statistical means of estimation should be considered as rather mixed: the F1/F2 values estimated from non-normalised data neither increased the comparability of formant plots across different speaker groups nor did they help to eliminate the vowel space shrinkage in recordings with a particularly poor spectral balance specification: the  $A_t$  values remained extremely similar across raw and estimated formant data (see Figure 5). In contrast, these vowel space characteristics were much less of an issue in Lobanov-transformed data where the estimates also more closely reflected the means of F1 and F2 (see Figure 6).

#### 5.4. Additional remarks on recording quality

Before moving on to making recommendations to future sociophonetic research involving real-time data, some additional remarks need to address the origins of the spectral imbalance in our data, preliminarily introduced in 2.3. Given previous research (see 1.2), we expected to find some spectral variability in our data resulting mainly from an occasional proximity effect. However, the distribution of spectral energy in our 1970s-samples suggested that the proximity effect alone could not have contributed to the reduced amount of energy available at higher frequencies. There was also some variation among individual recordings of the Ronald Macaulay’s set where the exact same recording equipment, set-up and procedure were used throughout the multiple recording sessions (Ronald Macaulay, personal communication). We consulted with the School of Scottish Studies Sound Archives (University of Edinburgh) who hold all of the Macaulay’s original reel-to-reel tapes and digitised them for our project, and with a colleague from LANCHART Centre in Copenhagen. A detailed inspection led to the conjecture that the above spectral issues may have been caused by the digitisation procedure itself; the tapes may have been slightly out of kilter (i.e. the tape and the tape head were misaligned) while the digitisation took place (Gert Foget Hansen from LANCHART, personal communication). Additionally, some dust on the playback head could have also led to the reduction of high frequencies. Fixing this problem required a technician to adjust the azimuth of the tape head on the playback machine to match the tape, and to clean the tape head.

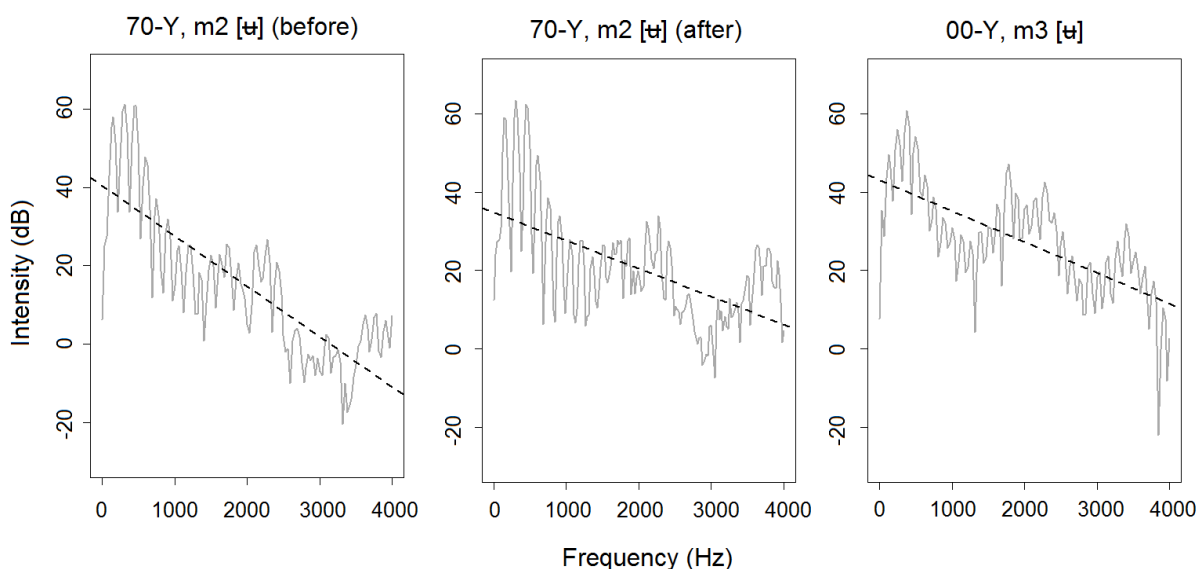


Figure 8. DFT-spectra of two [ʉ]-tokens taken from the midpoint of the vowel in ‘good’ produced by the speaker 70-Y, m2 (left: before the re-digitisation; right: after the re-digitisation). Dashed lines demonstrate the spectral tilt (for more detail on these spectra see Figure 2).

A re-digitisation of all affected tapes followed, with surprisingly impressive results. Figure 8 illustrates the difference in spectral tilts calculated for the same sound example (70-Y, m2; taken from Figure 2), prior to the re-digitisation and afterwards; and compares the 1970s data with the more recent recording made in the 2000s (00-Y, m3; taken from Figure

2). The spectral tilt of the re-digitised recording now more or less equals the unaffected recording, although the bass effect (higher energy in of F1 in comparison to F2 and F3) can still be attested.

This example demonstrates that even more technical factors of potential influence need to be taken into consideration when working with real-time data than initially hypothesised (see 1.2). Fortunately, a competent set-up of a digitisation is one of the factors over which sociophoneticians can exert more control, if aware of potential technical issues.

### ***5.5. Is there a sound approach to the inference of change?***

In this final section, we would like to offer some recommendations on how to approach potentially challenging and technically diverse real-time data in sociophonetic investigations, and to unveil the beauty in a beast.

First of all, sociophonetic research needs to show awareness of the technically introduced issues affecting both F1 and F2, by assessing the technical quality of the data with respect to at least SNR and spectral slope, before formants are measured (see 2.3). If only one of the two issues is present, the irrelevant factor can be ignored since their effects are independent of each other. If spectral imbalance is ascertained in the data, a potential redigitisation of the original tapes might help alleviate, if not completely extinguish, the problem. It is advisable to keep track of such technical information in a meta-data file for each recording of a sociolinguistic corpus, along with the information about speakers, interviewers and the recording situation.

Reliance on a single formant normalisation as a post-hoc method seems contraindicated. In our data, the transformation after Nearey retained a minimal amount of information about the poor SNR and the negative spectral balance of recordings in comparison to Bark, while the Watt-Fabricius method created formant values most affected by the technical issues. Following De Decker and Nycz (2010:54) who argue against using formant normalisations which rely on F3-measurements if the data cannot be assumed technically impeccable, we would like to also advise against applying the Watt-Fabricius transformation in such cases. Given that this normalisation procedure relies heavily on F1/F2 measurements of the corner vowels /i/ and /a/ whose values have been shown to be particularly affected by the two technical issues studied here, many carry-over effects are likely to obscure the patterns within a vowel system. The only transformation that we could show to reliably remove the technical influences of recording quality was Lobanov. However, it also removed sociolinguistically meaningful differences in the dataset: none of the relevant group comparisons showed significant results (see also Adank et al. 2004, Disner 1980).

In comparison to the above normalisation methods, the distance measure  $d_u$  showed little influence by the technical issues, similar to the Lobanov transformation. The  $d_u$ -dataset was, however, superior to the Lobanov-transformed data from the sociophonetic point of

view since only  $d_u$  clearly unveiled some meaningful effects involving /ʌ/-lowering and potential backing that we expected to find given existing evidence from independent research (Scobbie 2011; Scobbie, Lawson and Stuart-Smith 2012). Despite its advantage in context of a technically diverse corpus, a distance metric like the  $d_u$  measure comes with its own limitations. First, a metric of this type is not appropriate for investigations with an interest in the visualisation of the overall vowel space and dispersion. Stability of the reference vowels might be the second issue that would obscure the patterns of variation of change in the vowels in the centre of investigation. Moreover, direct comparisons of similar sound changes across different accents of English can be conceivably difficult if the reference vowels have different qualities.

Our general recommendation, then, is to combine the two approaches. Using Lobanov-normalised data, formant plots can be created to visualise the vowel space, its dispersion and the relationships between vowel categories. Notably, patterns of /ʌ/-lowering and retraction are somewhat visible in Figure 6, even though they did not reach significance levels in statistical tests. Subsequently, distance metrics for the variable(s) in question may help and narrow down the tendency of the change.

### **5.6. Summary and outlook**

The present study set out to examine whether or not technical quality differences present in a sociolinguistic real-time corpus might have an impact on the values of F1 and/or F2 measured by a standard LPC-algorithm implemented in Praat (Anderson 1978; Boersma and Weenink 2013). Sufficient evidence supported the idea of some technically introduced artefacts of F1/F2, which derived from noisy and spectrally compromised recordings. Although the magnitude of the F1/F2 deviations seemed rather small in these data, the technical effects could potentially accumulate, given the independence of SNR and spectral balance. Moreover, we worked with a relatively small dataset comprising of 24 recordings; a larger dataset (with more power) is likely to lend higher relevance to the effects that were just trending towards significance in our data.

Based on the evidence provided in our study, we recommend that sociophonetic investigations of real-time data consider some possible technical effects before a meaningful analysis of the sociolinguistic variation is conducted. The researchers may not always have control over the various factors influencing the technical quality of spoken data during a recording session (and there always might be more, yet unknown, external factors of influence that need to be taken into consideration, as the digitisation issues showed us in the present study). However, we suggest that a preliminary acoustic analysis of SNR and spectral balance properties of recordings should suffice to give the researchers an appreciation of potential technical interferences and a post-hoc control over the arising issues.

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