1 2	A statistical model for the prediction of wind-speed probabilities in the atmospheric surface layer
3	G. C. Efthimiou ^{1a} , D. Hertwig ^b , S. Andronopoulos ^a , J.G. Bartzis ^c , O. Coceal ^d
4 5	^a Environmental Research Laboratory, INRASTES, NCSR Demokritos, Patriarchou Grigoriou & Neapoleos Str., 15310, Aghia Paraskevi, Greece.
6	^b Department of Meteorology, University of Reading, Reading, UK.
7 8	^c University of Western Macedonia, Dept. of Mechanical Engineering, Sialvera & Bakola Str., 50100, Kozani, Greece.
9 10	^d National Centre for Atmospheric Science, Department of Meteorology, University of Reading, Reading, UK.

11 Abstract

Wind fields in the atmospheric surface layer (ASL) are highly three-dimensional and 12 characterized by strong spatial and temporal variability. For various applications such as 13 wind comfort assessments and structural design, an understanding of potentially hazardous 14 15 wind extremes is important. Statistical models are designed to facilitate conclusions about the occurrence probability of wind speeds based on the knowledge of low-order flow statistics. 16 Being particularly interested in the upper tail regions we show that the statistical behavior of 17 18 near-surface wind speeds is adequately represented by the Beta distribution. By using the properties of the Beta probability density function in combination with a model for estimating 19 extreme values based on readily available turbulence statistics, it is demonstrated that this 20 novel modelling approach reliably predicts the upper margins of encountered wind speeds. 21 The model's basic parameter is derived from three substantially different calibrating datasets 22 of flow in the ASL originating from boundary-layer wind-tunnel measurements and direct 23 24 numerical simulation. Evaluating the model based on independent field observations of nearsurface wind speeds showed a high level of agreement between the statistically modelled 25 26 horizontal wind speeds and measurements. The results show that, based on the knowledge of only a few simple flow statistics (mean wind speed, wind speed fluctuations and integral time 27 scales), the occurrence probability of velocity magnitudes at arbitrary flow locations in the 28 ASL can be estimated with a high degree of confidence. 29

30 Keywords

Atmospheric surface layer; Beta distribution; Direct numerical simulation; Extreme wind
 speeds; Field experiment; Probability density function; Weibull distribution; Wind tunnel

33 **1. Introduction**

¹ Corresponding author. Tel: +30 2106503405. E-mail address: gefthimiou@ipta.demokritos.gr (G. C. Efthimiou)

Knowledge of the probability density function (p.d.f.) of wind speed in the atmospheric boundary layer (ABL) is necessary for many applications such as the estimation of the wind energy potential at a particular location (Sarkar et al., 2011) or wind comfort and safety studies (e.g. Janssen et al., 2014). Furthermore, safety considerations in the design of structures at exposed locations (e.g. bridges, radio masts or wind turbines) require a reliable assessment of the maximum expected wind speeds (Brabson and Palutikof 2000).

40 For the prediction of wind gusts, i.e. sudden, brief increases of local wind speeds, various methodologies have been proposed in the past. Zhang et al. (2013), for example, performed 41 an analysis of the characteristics of 1-Hz wind-speed data continuously sampled in the air 42 layer below 2 m. They proposed an empirical model to predict the fluctuating wind gusts of 43 the streamwise velocity based on friction velocity, mean wind speed and standard deviation at 44 2 m. Sallis et al. (2011) used a machine-learning approach to determine meaningful and 45 robust results of wind gusts and proposed an algorithm for application to real-time climate 46 data. Brasseur (2001) proposed a new wind gust estimate method where the determination of 47 gusts is fully based on physical considerations. The proposed approach assumes that surface 48 49 gusts result from the deflection of air parcels in the upper levels of the boundary layer, which are brought down by turbulent eddies. The method takes into account the mean wind and the 50 turbulent structure of the atmosphere. 51

52 Over the past years, research activity in the area of wind-speed distribution modelling has 53 increased considerably. For the prediction of wind-speed distributions statistical models that 54 provide information about the local occurrence probabilities at a certain site are preferably 55 employed. For this task, the choice of a suitable p.d.f. is crucial. A number of previous 56 studies compared statistical distributions with measurements in order to examine how well 57 the p.d.f.s describe the statistical properties of the measured wind speed. An overview of 58 recent studies is presented in Table 1.

59	Table 1. Overview of recent studies (in chronological order) that use different statistical distributions to assess
60	wind speed occurrence probabilities.

Publication	Distributions	Averaging time interval of data analyzed
Donk et al. (2005)	Weibull	1 h
Carta et al. (2009)	Generalized Gamma, Gamma, Weibull, singly truncated from below normal, two components mixture Weibull, Rayleigh, beta, inverse Gaussian, lognormal	1 h
He et al. (2010)	Weibull	1 h
Morgan et al. (2011)	Gamma, lognormal, Rayleigh, log Pearson type III, Generalized Rayleigh, Generalized Gamma, Pearson type III, Weibull, Generalized normal, Wakeby,	10 min

	Kappa, bimodal Weibull mixture	
Waewsak et al. (2011)	Weibull	10 min
Aidan (2011)	Normal, Gamma, Weibull, Rayleigh	1 month
Sarkar et al. (2011)	Weibull, extreme value distribution of type I (Gumbel)	1 h
Odo et al. (2012)	Weibull	1 d
Kollu et al. (2012)	Weibull-extreme value distribution (GEV), Weibull-lognormal, GEV- lognormal	10 min
Masseran et al. (2013)	Lognormal, Weibull, Rayleigh, exponential, Gamma, inverse Gaussian, Burr, inverse Gamma	1 h
Datta and Datta (2013)	Weibull, exponentiated Weibull	1 a
Nemeş (2013)	Weibull	1 h
Indhumathy et al. (2014)	Weibull	1 h
Kidmo et al. (2015)	Weibull	1 h
Petković (2015)	Weibull	-
Men et al. (2016)	Gauss	30 min
Karthikeya et al. (2016)	Weibull	10 min
Carneiro et al. (2016)	Weibull	10 min

Table 1 shows that a large number of different p.d.f.s were previously compared with wind speed data, with the Weibull distribution overall being the most popular choice.

In order to assess extreme wind speeds, extreme value theory can be used (e.g. Palutikof et al. 64 1999; Holmes and Moriarty 1999; Simiu et al. 2001). However, in this case, the successful 65 modelling of the upper tail can often lead to an inadequate representation of the main part of 66 the wind-speed distribution. Most of the studies listed in Table 1 used 1-h or 10-min 67 averages. Steinkohl et al. (2010) analyzed the wind-speed measurements on a finer time scale 68 in the so-called micrometeorological range. Their dataset consisted of observations measured 69 in the ABL on two different days with a sampling frequency of 1 Hz. They focused on the 70 modelling of the tail of the wind-speed distribution by using the 'peaks-over threshold' 71 72 approach of extreme value theory.

Other studies use nonparametric estimators of wind speed. Rozas-Larraondo et al. (2014), for
 example, studied a new method based on nonparametric multivariate locally weighted

75 regression for improving wind speeds forecast by a numerical weather prediction model. 76 Wind direction data were used to build different regression models, as a way of accounting for the effect of surrounding topography. Recently, D'Amico et al. (2014) presented a new 77 nonparametric model to predict wind speeds based on semi-Markov chains. They found the 78 model to be able to reproduce the statistical behavior of wind speeds accurately for different 79 80 time scales when used as a forecast tool. Francisco-Fernández and Quintela-del-Río (2013) applied nonparametric curve estimation methods to analyze time series of wind speeds, 81 focusing on extreme events exceeding a chosen threshold. Nonparametric methods to directly 82 estimate quantities such as the probability of exceedance, the quantiles or return levels or the 83 return periods were proposed. Moreover, bootstrap techniques were used to develop 84 pointwise and simultaneous confidence intervals for these functions. 85

86 **1.1. Aim of this study**

In this study we aim at demonstrating that occurrence probabilities of wind speeds in the atmospheric surface layer (ASL) can be estimated based on the knowledge of low-order flow statistics readily available from field measurements and the choice of a p.d.f. that has a finite range. In quantitative terms we aim to predict the horizontal wind speed V within the time interval $\Delta \tau$ that is encountered at an arbitrary location,

92
$$V(\Delta \tau) = \frac{1}{\Delta \tau} \int_{\Delta \tau} v(t) dt$$
(1)

93 where v(t) is the instantaneous wind speed. $V(\Delta \tau)$ is in meteorological terminology the $\Delta \tau$ 94 gust.

The time interval $\Delta \tau$ can signify a measurement time interval or an averaging period, as is the 95 case in this study. However, $\Delta \tau$ can be assigned further significance when associating it with 96 typical exposure times to certain wind speeds, e.g. based on the time an individual is expected 97 98 to stay at a particular location. While mean velocities in urban areas are typically low, turbulence levels can be significant and wind gusts can by far exceed the time-averaged wind 99 speeds encountered at street level, which can cause discomfort for pedestrians or result in 100 structural damage to buildings. The time scales associated with such gust episodes are as low 101 102 as times associated with typical pedestrian walking speeds (i.e. a few s).

103 The derivation of a new p.d.f. for $V(\Delta \tau)$ and the associated maximum expected wind speed 104 $V_{max}(\Delta \tau)$ will be the main effort in the present investigation. A suitable p.d.f. should satisfy 105 the following criteria:

- 106 1. It should describe all parts of the wind-speed distribution accurately.
- 107 2. It should have a finite upper extreme.
- 108 3. It should be applicable to a wide range of turbulent flows in the ASL.
- 109 One probability distribution that fulfils these criteria is the Beta distribution.

The methodology employed in this study is described in detail in the next section. The experimental and numerical wind speed databases are introduced in Sect. 3. Based on this calibrating data, in Sect. 4 the basic parameter of the proposed model is estimated. Finally, in Sect. 5 the methodology is validated based on independent field data.

114 **2.** Methodology

The present methodology is based on the studies by Bartzis et al. (2008) and Bartzis et al. (2015), in which the cumulative distribution function (c.d.f.) of scalar concentrations from point source releases of airborne materials was modelled. In the present study we apply the same methodology to the prediction of wind speed p.d.f.s and c.d.f.s.

For this, we make the hypothesis that the maximum wind speed that can be encountered in the ASL takes a finite value and that this maximum wind speed $V_{max}(\Delta \tau)$ can be predicted by the Bartzis et al. (2008) model. We present a detailed justification for the applicability of the model in the following.

123 **2.1** Probability density function selection and parameterization

124 It is assumed that the p.d.f. for the time-averaged horizontal wind speed $V(\Delta \tau)$ at a certain 125 location is given by the Beta distribution (e.g. Gupta and Nadarajah, 2004)

126
$$pdf(x) \propto x^{\alpha-1} (1-x)^{\xi-1}; \qquad 0 \le x \le 1.$$
 (2a)

127 In the case of our study x is given by a normalized form

128
$$x = \frac{V(\Delta \tau)}{V_{\text{max}}(\Delta \tau)},$$
 (2b)

129 and hence x ranges from 0 to 1.

130 The exponents α and ξ are estimated from the wind speed mean, variance and the maximum 131 value based on the general relationships for Beta distributions,

132
$$\alpha = \frac{1}{1+\eta} \left(\frac{\eta}{I} - 1 \right), \tag{3a}$$

133
$$\xi = \eta \alpha \tag{3b}$$

134
$$\eta = \frac{V_{\max}(\Delta \tau) - \overline{V}}{\overline{V}}$$
 (3c)

135 where $V_{max}(\Delta \tau)$ is the maximum time-averaged wind speed in the interval $\Delta \tau$ or in 136 meteorological terminology the maximum $\Delta \tau$ gust in a large ensemble or a long time series. 137 We obtain $V_{max}(\Delta \tau)$ from the model proposed by Bartzis et al. (2008) (see Sect. 2.2), \overline{V} is the 138 mean wind speed and *I* is the wind speed fluctuation intensity given by

139
$$I = \frac{\sigma_V^2}{\overline{V}^2}$$
(4a)

140 with

141
$$\sigma_V^2 = \overline{V'^2} , \qquad (4b)$$

142 where σ_V^2 is the variance and V' is the fluctuation, which are quantities that are routinely 143 available from different types of atmospheric flow models by solving the relevant equations 144 (e.g. Hertwig et al., 2012; Koutsourakis et al., 2012) or from experimental measurements. 145 Hence, in order to use the model the extreme value $V_{max}(\Delta \tau)$ needs to be estimated. This is 146 described below.

147 2.2 Extreme value analysis

148 When adopting any finite range p.d.f. for the wind speed $V(\Delta \tau)$, the ability to estimate the 149 extreme value $V_{max}(\Delta \tau)$ forms a prerequisite. It should be noted that the experimental 150 maximum $V_{max}^{meas}(\Delta \tau)$ cannot be a priori expected to be the "true" extreme value as a direct 151 consequence of statistical uncertainties associated with limited measuring (or simulation) 152 times. For this reason we use the theoretical approach proposed by Bartzis et al. (2008) in 153 order to approximate the expected ("true") $V_{max}(\Delta \tau)$, which is modelled by

154
$$V_{\max}(\Delta \tau) = \overline{V} \left[1 + b \left(\frac{\Delta \tau}{T_V} \right)^{-\nu} I \right]$$
 (5)

where T_V is the wind speed integral time scale derived from the wind speed autocorrelation function $R_V(\tau)$ via

157
$$T_V = \int_0^\infty R_V(\tau) d\tau$$
 (6a)

158 and $R_V(\tau)$ is defined as

159
$$R_V(\tau) = \frac{\overline{V'(t)V'(t+\tau)}}{\overline{V'^2}}$$
. (6b)

Eq. 5 was developed initially for the estimation of maximum concentrations of airborne pollutants released from point sources. The theoretical background for the development of Eq. 5 is based on the application of the following equation for a stationary time series of infinite length:

164
$$\frac{V_{\max}(\Delta \tau)}{V_{\max}(\Delta T)} = \left(\frac{\Delta \tau}{\Delta T}\right)^{-n}$$
(7)

165 Eq. 7 is used for the calculation of the maximum expected wind speed in a time interval $\Delta \tau$ 166 when the maximum wind speed in a time interval ΔT is known. It follows a similar functional 167 form as proposed by Bartzis et al. (2008) for concentrations. Following the same argument 168 presented in Bartzis et al. (2008), here referring to wind speed instead of concentration, Eq. 7 169 is applied in deriving Eq. 5 based on the following assumptions:

- 170 (1) The maximum wind speed $V_{max}(\Delta T)$ tends to the mean wind speed \overline{V} as the time 171 interval ΔT increases.
- 172 (2) The time interval by which the wind speed $V_{max}(\Delta T)$ approximates \overline{V} is analogous to 173 the integral time scale of the wind speed T_V (Eq. 6a).
- 174 (3) When the time interval $\Delta \tau$ increases, the wind speed $V_{max}(\Delta \tau)$ approximates zero 175 according to Eq. 7, while it should tend to \overline{V} . This accounts for the additional term of 176 unity in Eq. 5.
- 177 (4) The ratio $V_{\text{max}}(\Delta T)/\overline{V}$ depends on the fluctuation intensity *I* (Eq. 4a).

Under these considerations the parameters that determine the extreme wind speed at a certain location are the fluctuation intensity I and the integral time scale T_V , together with the constants b and v.

The relationship presented in Eq. 5 was previously used successfully to predict maximum 181 time-averaged pollutant concentrations from near-ground emission sources based on 182 numerical results from simulations performed with computational fluid dynamics (CFD) 183 184 models (Effhimiou and Bartzis 2011, 2014; Effhimiou et al. 2011a, 2011b, 2015). The parameters b and v in Eq. 5 can be derived empirically and typically exhibit a wide range of 185 values as demonstrated in previous studies. This is a result of the combination of limitations 186 187 of the model, experimental errors, insufficient stationarity of the time series and the finite duration of the analyzed signal used to derive these values. Previous studies on the dispersion 188 of airborne material in atmospheric flows suggested indicative values of b = 1.5 and v = 0.3. 189

- 190 Several previous studies of wind gusts in the ASL have demonstrated that local gusts scale with the standard deviation of wind speed observed at the site (e.g. Beljaars 1987; Kristensen 191 1991), a parameter which indirectly includes information of surface roughness characteristics 192 and effects of atmospheric stratification. This dependence is also included in the model 193 shown in Eq. 5, via the fluctuation intensity I. Further information about the temporal scale of 194 the phenomena is added by including a direct link to the local auto-correlation time scale T_V . 195 The rationale for adopting the Bartzis et al. (2008) concentration model for the prediction of 196 wind speed extremes (Eq. 5) is further based on the following: 197
- The wind speed and the concentration are scalars, real numbers and take positive values.
- There is relation between the wind and the concentration which is expressed through
 the Schmidt number.
- 3. It is expected that both variables have finite extreme values in the ABL.
- 4. The assumptions that were mentioned before for the construction of Bartzis et al.
 (2008) model are also considered to be valid for the wind speed.

The last point is examined by analyzing sample data. For this purpose a random wind speed time series was selected from one of the experimental test cases used in this study (BL3-0 case described in Sect. 3.1.1.), which represents measurements in a rough-wall boundary layer modelled in the wind tunnel.

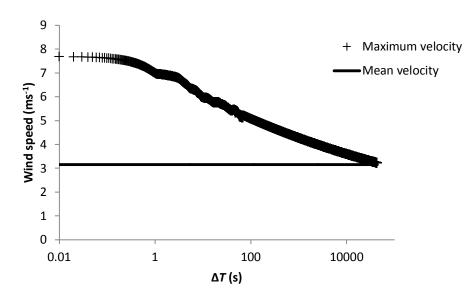
209 The first assumption is: "The maximum wind speed $V_{max}(\Delta T)$ tends to be equal to \overline{V} as the

210 time interval ΔT increases." The maximum time-averaged wind speed $V_{max}(\Delta T)$ of the

- random wind speed time series is plotted against ΔT in Fig. 1. Also the mean velocity is
- 212 presented. The horizontal axis is plotted in logarithmic scale. It is clear that $V_{max}(\Delta T)$

213 approximates \overline{V} with the increase of ΔT .

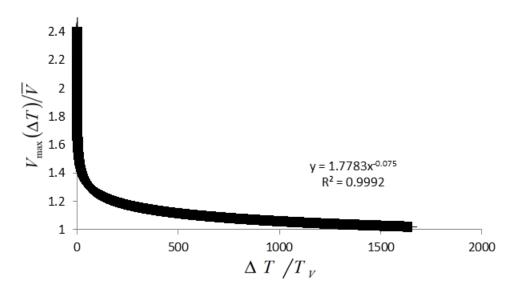




215

Fig 1 The maximum time-averaged wind speed and the mean wind speed are plotted versus the time interval ΔT . The results correspond to a random wind speed time series of the BL3-0 case.

The second assumption is: "The time interval by which the wind speed $V_{max}(\Delta T)$ approximates \overline{V} is analogous to the integral time scale of the wind speed T_V (Eq. 6a)." The ratio $V_{max}(\Delta T)/\overline{V}$ of the random wind speed time series is plotted against $\Delta T/T_V$ in Fig. 2. There is a clear correlation ($R^2 = 0.9992$) and a power-law function fits the data very well.



224 Fig 2. The ratio $V_{\text{max}}(\Delta T)/\overline{V}$ is plotted versus the ratio $\Delta T/T_V$ for the same dataset as in Fig. 1.

The third assumption is: "When the time interval $\Delta \tau$ increases, the wind speed $V_{max}(\Delta \tau)$ approximates zero, while it should tend to \overline{V} ." This is illustrated in Fig. 3.

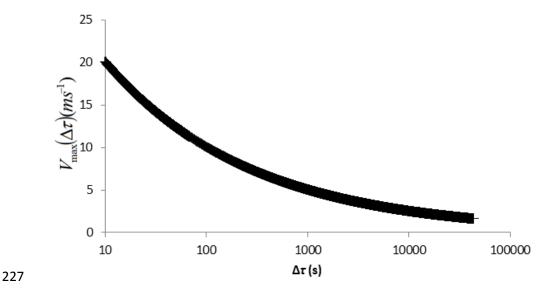


Fig 3 The maximum time-averaged wind speed $V_{max}(\Delta \tau)$ is plotted versus the time interval $\Delta \tau$ for the same dataset as in Fig. 1.

The fourth assumption is: "The ratio $V_{\text{max}}(\Delta T)/\overline{V}$ depends on the fluctuation intensity *I* (Eq. 4a)." In order to substantiate this assumption based on the data used in this study, the entire set of experimental wind speed time series of the wind-tunnel boundary-layer flow case BL3-

233 0 were analyzed. The ratio $V_{\text{max}}(\Delta T)/\overline{V}$ is plotted against the fluctuation intensity I in Fig. 4,

clearly illustrating a strong linear relationship ($R^2 = 0.947$).

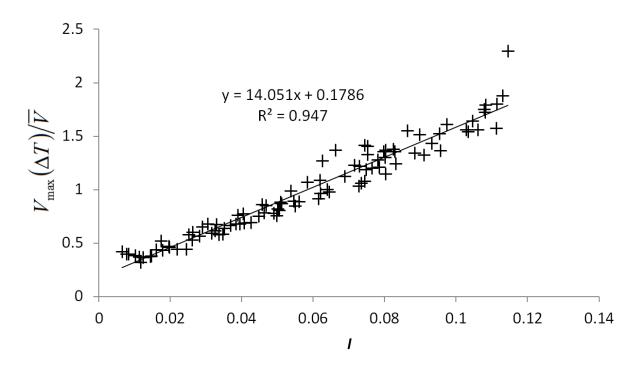


Fig 4 The ratio $V_{\text{max}}(\Delta T)/\overline{V}$ is plotted versus the fluctuation intensity *I* for all horizontal wind speed time series available from the rough-wall boundary-layer flow experiment in the wind tunnel.

Finally, a further test is performed for the applicability of the method. $V_{max}(\Delta \tau)$, \overline{V} , *I* and T_V used in Eq. 5 can be calculated from the experimental wind speed time series. In this case, $\Delta \tau$ denotes the measurement time interval. The indicative value of the parameter *v* in Eq. 5 is 0.3 if the equation is used to estimate peak concentrations. If we assume the same value to be valid for wind speed data, Eq. 5 can be rewritten as follows:

243
$$\frac{V_{\max}(\Delta\tau)}{\overline{V}} - 1 = bI\left(\frac{\Delta\tau}{T_V}\right)^{-0.3}$$
(8)

244 The form of Eq. 8 is equivalent to the linear equation y = b x where:

245
$$y = \frac{V_{\text{max}}(\Delta \tau)}{\overline{V}} - 1$$
(9a)

246
$$x = I \left(\frac{\Delta \tau}{T_v}\right)^{-0.3}$$
(9b)

If Eq. 5 is valid for wind speed then y should clearly correlate with x. y and x are plotted in Fig. 5. As for Fig. 4, the data points in Fig. 5 are from all the horizontal wind speed time series of the laboratory experiment (BL3-0 case). The correlation coefficient R^2 is 0.87, which indicates that there is a significant linear relation. Furthermore it is noteworthy that the parameter b takes a value of 1.85 which is close to the indicative value of 1.5 that had earlier been determined for concentration data.

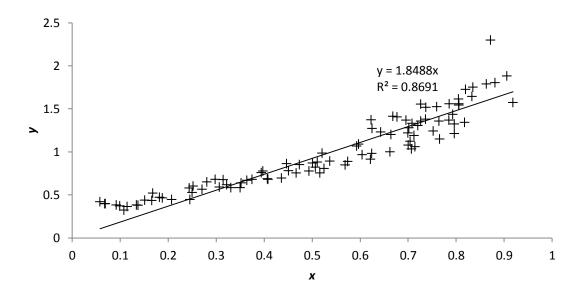


Fig 5 Examination of the validity of Eq. 5 for wind speed data based on all horizontal wind speed time series available from the rough-wall boundary-layer flow experiment in the wind tunnel.

However, since we are interested in the p.d.f.'s upper bound we go beyond the indicative maximum values measured in the experiments and instead focus on the extreme value $V_{max}(\Delta \tau)$ reached within a time interval of infinite duration. It is noted that naturally such an extreme value cannot be verified experimentally. When comparing the measured peak value $V_{max}^{meas}(\Delta \tau)$ with the expected extreme value $V_{max}(\Delta \tau)$ for a specific sensor location, in theory the relation $V_{max}^{meas}(\Delta \tau) \leq V_{max}(\Delta \tau)$ will always hold true due to the ultimately finite length of the measured signal.

It is proposed that the extreme value $V_{max}(\Delta \tau)$ can be approximated based on Eq. (5) given the values of the parameters *b* and *v*. The present strategy is to fix the value of v = 0.3 (the indicative value from Bartzis et al., 2008) and allow the *b* parameter to be estimated from suitable calibrating data. This reduces the problem of estimating $V_{max}(\Delta \tau)$ to estimating a single parameter, *b*.

Bartzis et al. (2008) have presented a method to estimate the parameter v, where $(V_{max}(\Delta \tau)/\overline{V}$ 268) - 1 is plotted versus $\Delta \tau$. $V_{max}(\Delta \tau)$ is the experimental maximum time-averaged wind speed 269 which changes at every $\Delta \tau$. In other words at every $\Delta \tau$ a new wind speed time series is 270 constructed and the maximum wind speed is calculated. For the following test $\Delta \tau$ ranges from 271 0.01 to 45000 s. V_{mean} is the experimental mean wind speed for $\Delta \tau$ equal to the measured time 272 interval i.e. the original time series that is measured by the instrument. For the following test 273 $\Delta \tau = 0.01$ s. For this test we selected a random time series from the BL3-0 wind-tunnel 274 experiment. Results are presented in Fig. 6. The equation $y = \alpha^*(x^{-\beta})$ was fitted to the data, 275 where y is $(V_{max}(\Delta \tau)/\overline{V})$ - 1 and x is $\Delta \tau$. In our case, the parameter β is the desired v 276 277 parameter. For this dataset β was found to be equal to 0.403, which is very close to the previously determined value of 0.3. Uncertainties arising after fixing the parameter v are 278 reflected in the value of b for each particular location. 279

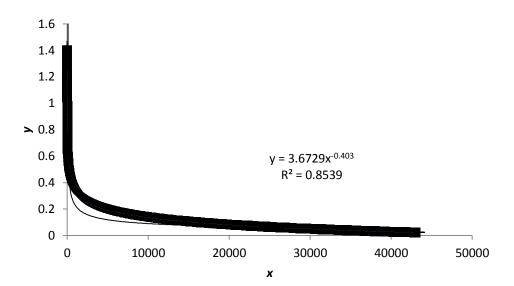


Fig 6 $(V_{max}(\Delta \tau)/V)$ - 1 is plotted versus the time interval $\Delta \tau$. The results correspond to a random wind speed time series of the BL3-0 case.

It is noted that for the parameters b and b^{meas} obtained from Eq. (5) using $V_{max}(\Delta \tau)$ and $V_{max}^{meas}(\Delta \tau)$ respectively, in theory $b^{meas} \le b$ always holds true since $V_{max}^{meas}(\Delta \tau) \le V_{max}(\Delta \tau)$.

With this approach we implicitly hypothesize the existence of a single value of *b* that can qualify as the upper bound of all b^{meas} values obtained at any location in the ASL flow at which the corresponding value of $V_{max}^{meas}(\Delta \tau)$ is detected.

- 288 To close the model, it remains to estimate the value of the parameter b in Eq. 5.
- 289 In summary there are two groups of equations.
- (1) The equations for the construction of the Beta distribution (Eqs. 2a, 2b, 3a, 3b and 3c).
- (2) The equation for the estimation of the maximum time-averaged wind speed (Eq. 5).
- The equations of the first group use the equation of the second group through the parameter η (Eq. 3c).

295 **2.3 Application of the method**

296 The methodology used in this study includes the following steps:

1. Estimation of the parameter *b* from an analysis of experimental and numerical datasets. In our study we use near-surface turbulent flow signals that are available from a dense sensor network, offering sufficient coverage of a diverse set of ASL flow scenarios, also including data measured within and above urban environments. The wind speed time series should be of high time resolution, statistically stationary and have a sufficiently long duration to ensure that relevant statistics are derived with high levels of statistical confidence.

- 303 2. Calculation of the mean wind speed, variance and time scale parameter T_V from each 304 measured wind speed time series from the calibrating data.
- 305 3. Identification of the peak value $V_{\max}^{meas}(\Delta \tau)$ and estimation of the corresponding local 306 parameter b^{meas} from Eq. (5) for each wind speed time series. In this study, the selected time 307 interval $\Delta \tau$ is equal to the time resolution of the experimental/simulation data.
- 4. Estimation of the parameters for the wind speed Beta p.d.f. as described by Eqs. (2a, b) and
- 309 (3a, b, c) for each location. The values of the required input variables \overline{V} , σ_V^2 and T_V are
- derived from the reference datasets (see point 2. above). The single missing parameter is the
- 311 coefficient b.
- 312 5. Derivation of a suitable value for *b*.
- 6. Evaluation of the accuracy and robustness of the single value for b derived from the calibrating data by testing the model based on independent wind speed measurements from field experiments.
- 316 It is emphasized that we look for a p.d.f. that produces a relatively simple but adequate 317 approximation. Thus, the present methodology is based on the assumption that if a theoretical 318 c.d.f. can reproduce the real threshold wind speeds for various probabilities then the p.d.f. 319 used is considered a good approximation.

320 3. Data and Test Cases

For the estimation of the parameter b and the construction of the statistical model we use data from wind-tunnel experiments and numerical simulations. The performance of the statistical model is evaluated in Sect. 5 based on hourly wind speed measurements taken at various field sites.

325 **3.1. Wind-tunnel experiments**

The first two flow datasets analyzed in this study to derive the model coefficient *b* stem from boundary-layer wind-tunnel measurements conducted at the Environmental Wind Tunnel Laboratory (EWTL) of the University of Hamburg. The datasets are part of the CEDVAL-LES reference database that offers time-averaged statistics as well as time-resolved data for different types of boundary-layer flow and dispersion scenarios under neutral stability conditions. The validation datasets are freely available and described in detail in Fischer et al. (2010).

333 CEDVAL-LES contains data for various levels of geometric complexity and surface 334 roughness characteristics. In the database, the term "complexity" refers to the configuration 335 of the flow scenario covered in the experiment, ranging from complexity 0 denoting simple 336 rough-wall boundary-layer flows, over flows around isolated obstacles or within obstacle 337 arrays (complexities 1 or 2, respectively) to flows in semi-idealized urban environments 338 (complexity 3) or realistic city layouts (complexity 4). Two of the available cases were 339 selected for this study: (1) boundary-layer flow over a very rough surface (CEDVAL-LES complexity 0; case reference: BL3-0) and (2) urban flow within and above a semi-idealized
city geometry (complexity 3; "Michel-Stadt" case reference: BL3-3). In both cases the flow
was physically modelled under a scale of 1:225 and point-wise velocity measurements were
conducted by means of 2D laser Doppler anemometry (LDA). The setup of the very rough
boundary-layer flow (BL3-0) was used as the inflow boundary layer for the semi-idealized
city case (BL3-3).

346 **3.1.1 Boundary layer over very rough surface**

With a power-law profile exponent of $\alpha = 0.27$ and a roughness length of $z_0 = 1.53$ m, the boundary-layer flow (complexity 0 case) shows roughness characteristics of flow above an urban environment. In this case, the buildings are not directly represented, but their aerodynamic effect on the approach flow boundary layer is physically modelled by means of floor-roughness elements.

At sufficient distance from the tunnel inlet and from the floor-roughness elements, the 352 353 boundary layer was verified to be horizontally homogeneous. Measurements are only taken above the blending height, where only the integrated effect of the surface roughness is 354 represented in the flow characteristics. The 2D-LDA was operated consecutively in two 355 measuring modes to acquire two components of the velocity vector at a time: the streamwise 356 (U) and spanwise (V) velocities (UV-mode) and the streamwise and vertical (W) velocities 357 (UW-mode). Data are available in terms of vertical profiles and horizontal transects 358 perpendicular to the mean inflow direction. All velocity data were scaled to a full-scale 359 reference height of $z_{ref} = 100$ m (444.44 mm model scale) with reference wind speeds U_{ref} 360 ranging between 4.75 and 6 m s⁻¹. With a measurement duration of 3 min (corresponding to 361 11 h full scale) per locations, the derived velocity statistics offered a high level of statistical 362 representativeness. 363

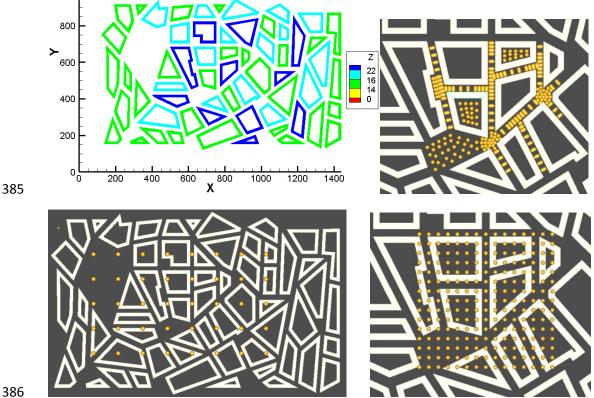
For the present study, wind speed time series with a full-scale resolution of $\Delta \tau = 0.01$ s were analyzed at 96 points for *UV* measurements. It should be noted that all time-series were resampled to the same time resolution.

367 **3.1.2** Flow in a semi-idealized city

The semi-idealized urban geometry includes typical features of Northern and Central European cities like courtyards, oblique road arrangements, squares and complex intersections. Three building heights are included in the model: 15, 18 and 24 m full-scale (see. Fig. 7 top left). All buildings had flat roofs.

Velocity measurements were conducted with the LDA in *UV*-mode, providing information about the horizontal winds within and above the city. The streamwise reference velocity at a height of 100 m was kept at 6 m/s and was monitored during each measurement run. The reference velocity was verified to be sufficiently high to guarantee Reynolds number independence of derived flow statistics within and above the urban canopy. Three groups of wind speed measurements were used in this study: (1) 40 vertical profiles distributed at various points throughout the city (Fig. 7, bottom left), (2) detailed measurements on dense 379 horizontal grids at heights of 2, 9 and 18 m full-scale within the main city area (340 m x 340 m), containing 383 measurement points per level (Fig. 7, top right); (3) measurements on a 380 coarse horizontal grid above the city centre at full-scale heights of 27.5 and 30.2 m, 381 containing 252 data points at each height (Fig. 7, bottom right). 382

The total number of signals analyzed is 2,158. Again the time series were resampled to a 383 resolution of 0.01 s. 384



391 **3.2 Direct numerical simulations**

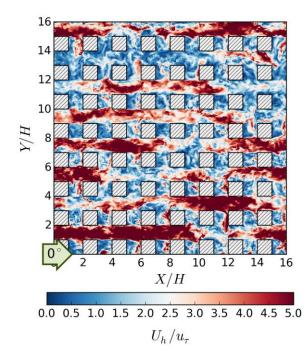
The second type of data analyzed in this study stems from simulations of urban flow fields in 392 an idealized urban roughness generated by direct numerical simulation (DNS). With the DNS 393 approach, turbulence is directly resolved down to the small dissipative eddy scales. In order 394 to facilitate this computationally, the flow is simulated at lower Reynolds numbers compared 395 to the ones typically encountered in the atmospheric boundary layer. The Reynolds numbers 396 397 that can be realized, however, are comparable to those typically achieved in boundary-layer wind-tunnel experiments as those described in the preceding sections. Compared to other 398 turbulence-resolving CFD approaches like for example large-eddy simulation (LES), the 399 accuracy of flow simulations with DNS is not affected by errors resulting from turbulence 400 401 modelling. Hence, DNS data, after appropriate accuracy checks, can be used as reference data similar to experimental measurements. The significant computational requirements involved 402

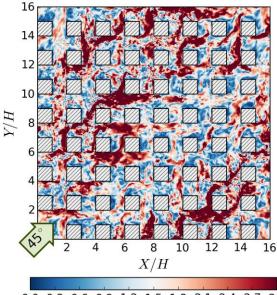
³⁸⁷ **Fig 7** Top left: the idealized city domain (dimensions given in full scale, dimensions X, Y and Z are in meters); top right: densely-spaced measurement locations within the urban canopy layer at elevations of 2 m, 9 m and 18 388 389 m; bottom left: vertical profile locations; bottom right: densely-spaced measurement locations above roof top at 390 heights of 27 m and 30 m.

in performing DNS, however, currently restrict the applicability of the technique to simpleflow scenarios at low Reynolds numbers.

The DNS code that had been used to generate the data analyzed in this study is the research 405 code CgLES developed at the University of Southampton specifically for performing 406 massively-parallel DNS and LES computations. The code is parallelized with the Message 407 Passing Interface (MPI) and a flexible multi-block mapping strategy is used to deal with 408 409 complex geometries. The Navier-Stokes equations are discretized using second-order central finite differences in space and a second-order Adams-Bashforth scheme in time based on the 410 pressure correction method. The Poisson equation for pressure is solved using a multigrid 411 method. A detailed description of the numerical techniques involved in the DNS as well as 412 examples of previous studies with the code can be found, e.g., in Yao et al. (2001), Branford 413 et al. (2011) or Coceal et al. (2006, 2007, 2014). 414

Flow simulations were conducted in a geometry comprised of 64 cubical obstacles of height 415 *H* that were set up in a regular array consisting of 8 rows of 8 obstacles. The computational 416 domain was of size 16H x 16H x 8H. Periodic boundary conditions were prescribed in 417 horizontal directions, free-slip conditions at the upper domain boundary and no-slip and 418 impermeability conditions at the bottom of the domain and at solid surfaces. The flow was 419 driven by a constant body force that resulted in a roughness Reynolds number of $Re_{\tau} = u_{\tau} H/v$ 420 = 500, where u_{τ} is the total wall friction velocity and v is the kinematic viscosity. Sensitivity 421 studies for this setup (Coceal et al. 2006, 2007) demonstrated that the selected resolution of 422 423 the uniform grid of H/32 is sufficient to adequately resolve the flow down to the dissipative scales. Data is available for two wind directions (0° and 45°) under neutral stratification. Fig. 424 8 shows snapshots of the instantaneous horizontal wind speed magnitudes at a height of Z =425 426 0.5*H* for both wind directions, illustrating the complexity of the canopy layer flow field.





0.0 0.3 0.6 0.9 1.2 1.5 1.8 2.1 2.4 2.7 3.0 $U_h \, / u_{ au}$

427

- 428 Fig 8 Plan views of the computational domain of the DNS including contours of instantaneous snapshots of the
- 429 non-dimensional horizontal velocity magnitude, U_h/u_t , at half the building height (Z = 0.5H). Left: 0° forcing
- 430 direction; right: 45° forcing direction.

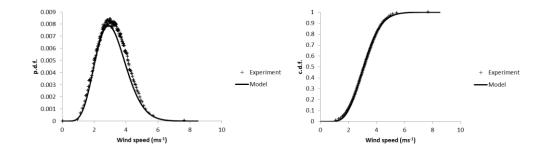
Wind speed time series at 896 locations overall (for both wind directions) were analyzed in 431 this study, each of them having a length of 140T, where T is the eddy turnover time defined 432 as $T = H/u_{\tau}$. As done with the wind-tunnel data, the DNS time series were resampled to a 433 non-dimensional time resolution of 0.01. The data locations are distributed in two horizontal 434 planes covering the entire simulation domain at two different heights: Z = 0.5H and Z = 1.5H. 435 Time series were extracted in the centre of each street and intersection. The DNS 436 computation were performed on a supercomputer and required a total spin-up time of 12 days 437 on 124 nodes for both the 0° and 45° runs, after which the simulations ran for10 days (0° 438 case) and 13 days (45° case) on 248 nodes to collect flow time series and statistics. 439

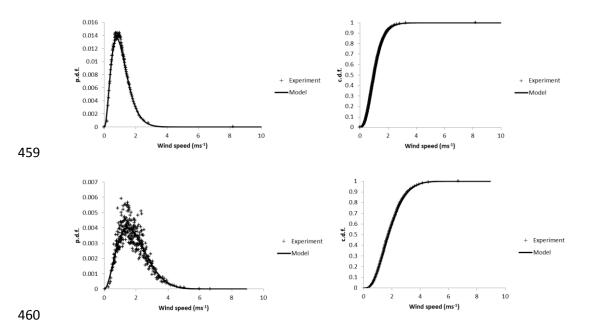
- 440
- 441 **4. Estimation of the** *b* **parameter**
- 442 4.1 Wind-tunnel flow fields
- 443 4.1.1 Boundary-layer flow

444 The autocorrelation time T_V is calculated from the autocorrelation function $R_V(\tau)$ on the 445 interval from 1.0 to 0. The wind speed time series are considered to be characterized by a 446 sufficiently high temporal resolution.

447 At each sensor location, an experimental peak $V_{\text{max}}^{\text{meas}}(\Delta \tau)$ is identified and a b^{meas} value is 448 estimated from Eq. 5 using $V_{\text{max}}^{\text{meas}}(\Delta \tau)$. The values of b^{meas} range from 1.5 to 7.2. The value of 449 5, however, is only exceeded three times out of 96 (i.e. 3.1%) and seems to point more to 450 'outliers' behavior.

- It is evident that if the proposed model is valid and a single value for *b* exists this value has to be greater than 5. This dataset and the two other datasets (discussed in detail in the following Sects. 4.1.2. and 4.2) clearly indicate that a value of b = 6 is appropriate for these flow scenarios.
- For a randomly selected sensor of the BL3-0 case, the Beta and the experimental p.d.f.s and c.d.f.s are plotted in Fig. 9 (first row). The agreement between the model curve and the experimental data points is very good.





461 Fig 9 Experimental and theoretical p.d.f.s and c.d.f.s using the Beta distribution for a randomly selected sensor
462 of the BL3-0 case (first row); for a randomly selected sensor of the BL3-3 case (second row) and for a randomly
463 selected sensor of the DNS experiment (third row). The parameter *b* is set to 6.

464 4.1.2 Semi-idealized city

As a next step, the robustness of using a value for b equal to 6 was examined by analyzing flow data acquired in the semi-idealized city wind-tunnel model.

467 As for the boundary-layer flow data, the b^{meas} is obtained based on the analysis of all 2,158 468 signals. The b^{meas} ranges from 0.63 to 7.1. Again, only at three of the available measurement 469 positions a b^{meas} value larger than 5 was obtained. The proposed value of *b* (equal to 6) is also 470 relatively close to the maximum value of 7.1.

As for the boundary layer flow case analyzed above, the agreement between the Beta and the
experimental p.d.f.s and c.d.f.s for an example time series from a randomly selected sensor
location is very high (Fig. 9, second row).

474 **4.2 Direct numerical simulations**

The analysis of the wind-tunnel data clearly supports the proposed methodological approach.
In a next step, the DNS data of flow in an idealized urban roughness is analyzed in a similar
manner.

As with the experimental data, the DNS dataset is first analyzed with regard to determining an appropriate value of the *b* parameter. For each sensor, a peak $V_{\text{max}}^{meas}(\Delta \tau)$ is identified from the simulation data and the corresponding b^{meas} value is estimated from Eq. (5). The b^{meas} range is from 1.1 to 5.56. Only 10 values are larger than 5 and the proposed *b* value (equal to

482 6) is relatively close to the maximum value 5.56.

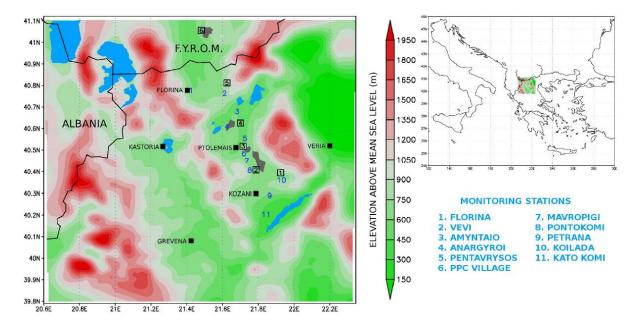
The Beta and the experimental p.d.f.s and c.d.f.s from time series at a randomly selected point of the DNS flow simulation is presented in Fig. 9 (third row) and again reveal a high level of agreements between the statistical model and the reference data.

486 It is worth highlighting that it is rather impressive that two completely different kinds of data 487 sources, the wind-tunnel measurements and the DNS, corresponding to very different near-488 surface flow scenarios show such a similar range of values determined for b^{meas} . This 489 provides strong support for the hypothesis that a single representative maximum value of the 490 *b* parameter can be derived, which could then be applied to various types of ASL flows.

491 5. Model evaluation based on field data

Based on the analysis of the three different calibrating datasets presented above, an upper 492 493 value of the parameter b equal to 6 was derived. However, the flow scenarios analyzed to derive this model parameter represent quite idealized cases of ABL flow (e.g. with respect to 494 the isothermal conditions/neutral stratification and stationarity of the flow). In order to 495 496 demonstrate the applicability of the model to real-world ABL flow scenarios, in a next step the model is evaluated based on independent field measurements. These consist of hourly-497 averaged in-situ wind-speed measurements from multiple ground-based sensors available 498 over the course of several months, which reflect the true variability of the natural atmospheric 499 boundary layer in terms of wind-speed trends (e.g. through the propagation of meso-scale 500 systems) and stratification effects. 501

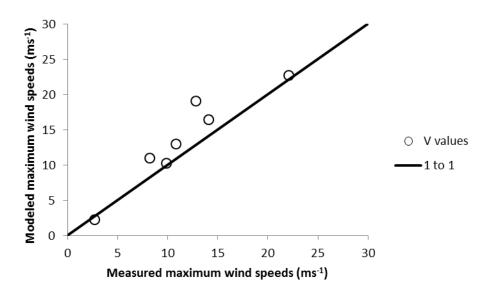
502 For this study, the Greek Public Power Corporation (<u>https://www.dei.gr/en</u>) provided hourly 503 wind-speed data for the period from 1 January 2012 to 31 August 2012 from 7 504 meteorological stations (Vevi, Florina, Koilada, PPC village, Pentabrysos, Petrana and 505 Pontokomi) located in the western part of Greece (see Fig. 10). All velocity sensors are 506 located in urban areas, the measurement height is 10 m above ground and the averaging 507 period of the signals is 1 h.



508

Fig 10 Topography of Western Macedonia, Greece, where the meteorological stations are located. The black
 numbered boxes indicate power plants, gray areas mines. The blue numbers indicate the location of
 meteorological measurement stations, with names listed on the right. Data available for this study are from the
 stations in Vevi, Florina, Koilada, PPC village, Pentabrysos, Petrana and Pontokomi.

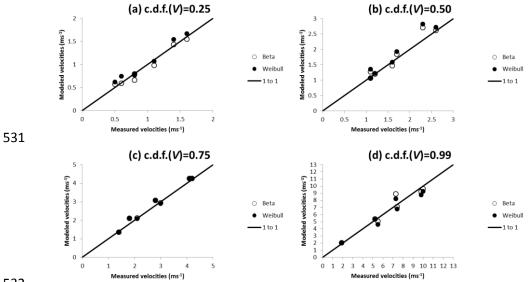
- 513 Initially a comparison is performed in terms of the theoretical value of $V_{max}(\Delta \tau)$ as derived
- from the Bartzis et al. (2008) model and the measured $V_{\text{max}}^{\text{meas}}(\Delta \tau)$ at all stations (see Fig. 11).
- 515 The model provides a success rate of 85.7% (only one value is below the 1:1 line), which
- supports the hypothesis that the proposed theoretical $V_{max}(\Delta \tau)$ can serve as an upper bound of
- 517 the corresponding measured $V_{\text{max}}^{\text{meas}}(\Delta \tau)$.



518

519 Fig 11 Modelled versus measured peak wind speeds at the 7 field measurement stations. The straight line 520 indicated the 1:1 relationship.

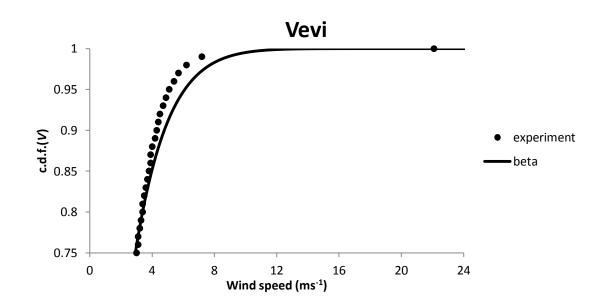
In the following analysis the performance of the Weibull distribution is also tested. This 521 distribution was fitted to the data and its parameters were calculated with the maximum 522 likelihood estimation. The results of the 99% threshold (c.d.f.(V)=0.99) are analyzed in order 523 to test the performance of the statistical model at the upper tail of the distribution. Based on 524 the analysis of the calibrating data discussed in the preceding section, the Beta distribution 525 was configured with a b value of 6. In Fig. 12 scatter plots are presented comparing the 526 calculated wind speed $V(\Delta \tau)$ from the Beta and Weibull distributions and the ones derived 527 from the field data corresponding to cumulative probabilities of c.d.f. (V) = 0.25, 0.50, 0.75528 and 0.99. The Beta distribution performs slightly better than the Weibull distribution in the 529 higher wind-speed range for probabilities of 0.25, 0.50 and 0.99. 530



532

Fig 12 Modelled (Beta and Weibull) versus measured wind speeds from the field experiment corresponding to (a) c.d.f.(V) = 0.25, (b) c.d.f.(V) = 0.50, (c) c.d.f.(V) = 0.75 and (d) c.d.f.(V) = 0.99.

The remaining critical question is how individual c.d.f.s derived from the Beta distribution compare with the counterparts from the field measurements at each location. In Fig. 13 the measured and modelled c.d.f.s for percentiles between 75th and 100th are shown for three stations (Vevi, PPC village and Pontokomi). At these and the other locations not shown here, the model presents good agreement with the field data.



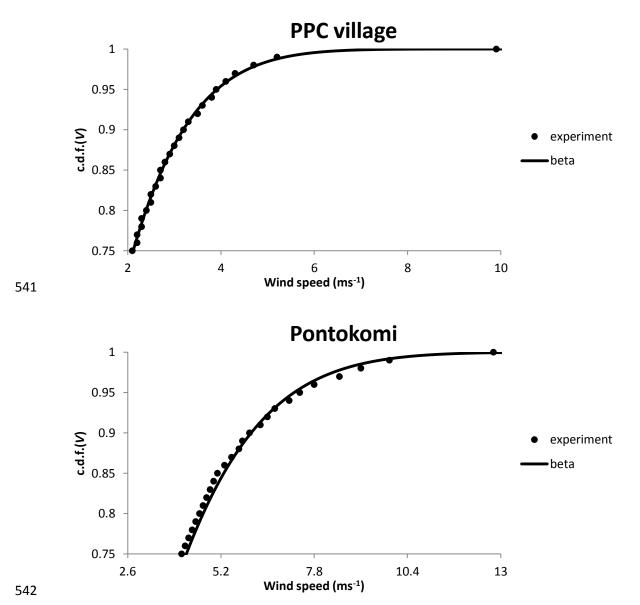


Fig 13 Modelled versus measured c.d.f. results for percentiles between the 75th and 100th for three meteorological stations of the Western Macedonia region of Greece.

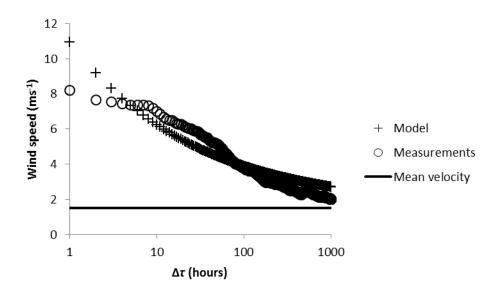
Concerning the atmospheric stability it should be noted that the wind tunnel and the DNS experiments modelled neutral conditions. However the present field experiment covers all the stability conditions and the Beta distribution performs very well under all conditions. Using a value of b = 6 when applying the model to independent datasets from field measurements shows that the value seems to be a good choice when dealing with ASL wind speeds.

550 5.1 Effect of the time interval $\Delta \tau$ on the performance of the model

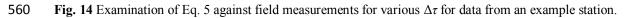
551 A further testing of the universal nature of the proposed Beta model is conducted by 552 repeating the above analysis for different $\Delta \tau$.

In Fig. 14 the peak time-averaged wind speeds based on Eq. 5 and the field experiment are plotted for different $\Delta \tau$ for an example station. The horizontal axis is presented using a logarithmic scale. For $\Delta \tau = 1$ h the model wind speed is higher than the experimental peak

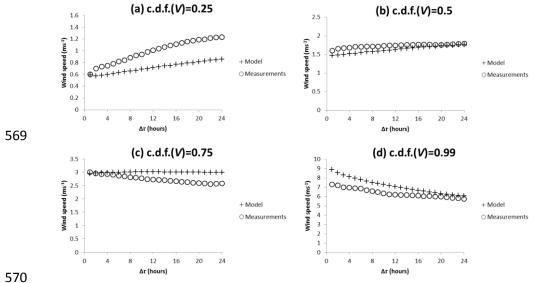
wind speed, as expected. For $\Delta \tau$ between 10 h and 100 h the model slightly underestimates 556 the experimental values, while overestimating again after 100 h. For 1000 h the $V_{max}(\Delta \tau)$ of 557 both model and experiment approximate the mean wind speed. 558



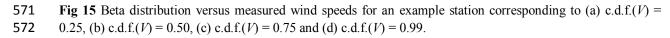




In Fig. 15 $V(\Delta \tau)$ obtained from the Beta distribution model and the values derived from one 561 of the field stations are presented. $\Delta \tau$ ranges from 1 h to 24 h. The results correspond to 562 cumulative probabilities of c.d.f.(V) = 0.25, 0.50, 0.75 and 0.99. For c.d.f.(V) = 0.25 and 0.5 563 the wind speed increases with $\Delta \tau$ while for 0.75 and 0.99 it decreases with $\Delta \tau$. This indicates 564 that with the increase of $\Delta \tau$ the distribution becomes more leptokurtic. The model shows the 565 same tendency as the experiment. For the specific station the largest deviation of the model 566 from the experiment is observed for c.d.f.(V) = 0.25 and for $\Delta \tau = 24$ h. On the other hand the 567 best performance of the model is observed for c.d.f.(V) = 0.5 over the entire range of $\Delta \tau$. 568







573 **5. Conclusions and Outlook**

574 By using the properties of the Beta p.d.f. in combination with a model for estimating extreme 575 values based on readily available turbulence statistics (Bartzis et al., 2008), this study 576 demonstrated that this novel modelling approach can reliably predict the upper margins of 577 wind speeds encountered in the ASL.

578 The problem itself is quite complex and adequate validation studies require extensive 579 experimental datasets. Such comprehensive validation efforts exceed the scope of a single 580 publication, but the work presented here represents a significant first step towards a thorough 581 testing of the proposed methodology.

The selected calibrating data for constructing the proposed model are representative of different scenarios of turbulent wind flow in the ASL: a rough boundary layer without buildings, a typical European urban micro-environment and an urban-like arrangement of cubical buildings. The sensor locations cover the boundary layer and the urban canopy-layer characteristics. The temporal resolution of the wind-speed signals covers a wide range of fluctuation intensities.

The performance of the model was successfully evaluated based on long-term independent field measurements (hourly averages), which cover the true variability of ABL flows in terms of wind-speed trends through the propagation of meso-scale systems and stratification effects. Concerning the atmospheric stability it should be noticed that the wind tunnel and the DNS experiments were conducted under neutral conditions. The field experiment, however, covered a wide range of stability conditions that can be encountered in natural ASLs and the Beta distribution performed very well under all conditions.

595 From the results obtained the following main conclusions are drawn:

596 1. The approximation of the statistical behavior of the abovementioned wind speed variability597 with a Beta distribution p.d.f. was shown to be satisfactory.

- 598 2. The important issue of the extreme value in the Beta distribution is properly addressed by599 an adaptation of the Bartzis et al. (2008) model.
- 600 3. The present work proposes b = 6 and v = 0.3 in Eq. 5.
- 4. With an increase of the averaging time interval the wind-speed distributions of the modeland experiment become more leptokurtic.
- The new model can broaden the capability of ensemble-averaged computational models such as Reynolds Averaged Navier Stokes–CFD models to estimate the wind-speed p.d.f. provided that reliable predictions of mean wind speeds, wind-speed fluctuations and integral time scales are available from these computations.
- 607

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- 611 https://www.mi.zmaw.de/index.php?id=6339). The authors would like to thank the Greek
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