

Error estimates for $\pi\pi$ scattering threshold parameters in Chiral Perturbation Theory to two loops.¹

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Abstract

Using the analysis of ChPT to two loops, we perform an error analysis of the threshold and low energy parameters, based on the uncertainties for the one loop low energy parameters and the resonance saturation mechanism. Different sets of one loop low energy constants have been considered. Thus, the predictive power of the effective field theory, is quantified on the basis of the present experimental uncertainties.

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Chiral perturbation theory (ChPT) is the effective pion field theory where the expansion parameter is $m^2/(\sqrt{4\pi}f)^2$, with m the pion mass and f the weak pion decay constant. At some finite order an increasing number of undetermined parameters is generated which can, so far, only be fixed by experimental data. Besides the theoretical uncertainty introduced by the finite number of terms considered in the expansion, there is an additional experimentally induced uncertainty. The low energy parameters inherit such an uncertainty, thus limiting in practice the predictive power of the effective field theory; if the errors at a given order, say $\mathcal{O}(m^{2n}/(\sqrt{4\pi}f)^{2n})$ are larger than the contributions of the next order $\mathcal{O}(m^{2n+2}/(\sqrt{4\pi}f)^{2n+2})$, the calculation of the latter becomes useless, unless more accurate experiments are performed.

To make our points quantitative, we consider $\pi\pi$ scattering as the prototype reaction to study the role played by chiral symmetry breaking and the validity of the chiral expansion within QCD at low energies. The scattering lengths and effective ranges for the isospin $I = 0, 1, 2$ have been computed in $SU(2)$ ChPT at tree level [1], one loop [2] and two loops [3, 4], in a power series expansion

$$a_{IJ} = a_{IJ}^{\text{tree}} + a_{IJ}^{1\text{loop}} + a_{IJ}^{2\text{loop}} + \dots \quad (1)$$

$$b_{IJ} = b_{IJ}^{\text{tree}} + b_{IJ}^{1\text{loop}} + b_{IJ}^{2\text{loop}} + \dots \quad (2)$$

At tree level, the number of parameters involved are two: the pion weak decay constant $f_\pi = 93.2$ MeV and the pion mass $m = 139.6$ MeV. For the purposes of the present

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discussion the experimental error bars in these parameters can effectively be taken to be zero. At the one loop level one has four, in principle undetermined, parameters $\bar{l}_1, \bar{l}_2, \bar{l}_3, \bar{l}_4$. At the two loop level the whole $SU(2)$ $\pi\pi$ amplitude information can be gathered into six coefficients, which in Ref. [3, 4] have been called b_1, \dots, b_6 . Ideally, these parameters could be extracted from a direct low energy analysis of $\pi\pi$ scattering experiments. The data are, however, too poor in the low energy region and several methods have been devised. Motivated by the success of the resonance saturation hypothesis at scales $\mu \sim 0.5 - 1$ GeV, at the one loop level [5], the authors of Ref. [4] suggest to estimate the bulk of the two loop corrections through resonance saturation. This approach requires the values for the one loop \bar{l} 's deduced from some clean source. In Ref. [4] two different sets of \bar{l} parameters have been considered: i) by taking the \bar{l}_1 and \bar{l}_2 parameters obtained from K_{l4} form factors using a dispersive one loop calculation for three flavors [6] (Set **I** in [4]), and ii) fixing \bar{l}_1 and \bar{l}_2 to the values deduced from D -wave $\pi\pi$ scattering lengths, at order $\mathcal{O}(p^6)$ (Set **II** in [4]). Unfortunately, no error analysis has been undertaken in Ref. [4]. On the other hand, the estimates based on K_{l4} analysis, can now be improved thanks to a recent study of this process at two loop accuracy [7]. Other possibility is to extract the one loop low energy constants, together with their errors, from Roy sum rules, saturated by the high energy behavior of $\pi\pi$ scattering [8]. In conjunction with resonance saturation, this procedure also yields an error analysis of threshold parameters [8].

An error analysis from the point of view of the predictive power of ChPT is missing, and it is the main subject of the present work. We assume that primary quantities, coming from a χ^2 -fit or direct experimental measurements, are Gauss distributed (with or without correlations), as we learn from elementary statistics. We propagate errors by means of Monte Carlo simulations, keeping always statistical correlations between all parameters entering in a given derived quantity. We use 10^4 samples, and to quote errors we always use a 68% confidence level around the central value. Since the outgoing threshold parameters distributions are not gaussians, we take in this way into account possible skewness in the distributions. By using the Monte Carlo method we avoid summing errors in quadratures, which would be incorrect for statistically correlated quantities, and we do not have to use any complicated covariance formula.

Let $a^{(n)}$ and $\Delta a^{(n)}$ be the n -loop central value and error of an observable. Thus, to be *predictive* and *convergent* at the n -loop level one ought to have the relation,

$$\Delta a^{(n)} \ll a^{(n+1)} \ll a^{(n)}. \quad (3)$$

In order to establish these necessary requirements from a quantitative point of view we will adopt the currently accepted scheme of assuming that resonance saturation yields the bulk of the two loop contribution at scales $0.5 - 1.0$ GeV. This induces a scale ambiguity in the two-loop part of the b -parameters, which we implement by providing the scale with an error,

$$\mu = 750 \pm 250 \text{ MeV}. \quad (4)$$

Moreover, since one does not expect resonance saturation to be exact, we provide these parameters with 100% uncertainty and take, in the notation of Ref. [4], the values

$$\begin{aligned} r_1 = -0.6 \times 10^{-4} \times (1 \pm 1) \quad r_2 = 1.3 \times 10^{-4} \times (1 \pm 1) \quad r_3 = -1.7 \times 10^{-4} \times (1 \pm 1) \\ r_4 = -1.0 \times 10^{-4} \times (1 \pm 1) \quad r_5 = 1.1 \times 10^{-4} \times (1 \pm 1) \quad r_6 = 0.3 \times 10^{-4} \times (1 \pm 1) \end{aligned} \quad (5)$$

at the scale μ given in Eq. (4). In addition to these numbers, one needs estimates of the one loop low energy parameters $\bar{l}_{1,2,3,4}$. Through all this work we fix

$$\bar{l}_3 = 2.9 \pm 2.4, \quad \bar{l}_4 = 4.4 \pm 0.3, \quad (6)$$

as determined from the study of $SU(3)$ breaking effects and the scalar form factor in Refs. [2] and [9] respectively. There is less consensus regarding the values of \bar{l}_1 and \bar{l}_2 , therefore we consider three possibilities. We give below the central values, errors and linear correlation coefficients for \bar{l}_1 and \bar{l}_2 in each case, in order to characterize the full statistical distributions, deduced from different primary distributions, which will be used later to calculate threshold and low energy parameters.

- Set **I** K_{l4} form factors: Ref. [7] allows to determine \bar{l}_1 and \bar{l}_2 from a two loop analysis of K_{l4} decays. The analysis is done in terms of the scale dependent $SU(3)$ low energy constants L_i^r at the ρ mass scale, $m_\rho = 770$ MeV, and central values for \bar{l}_1 and \bar{l}_2 are given. The relation between the $SU(2)$ and $SU(3)$ low energy constants is known at one loop [10] and thus the determination of the \bar{l}_1 and \bar{l}_2 parameters in [7] suffers from two loop $SU(3)$ uncertainties. From the remarks of that work a strong anti-correlation between \bar{l}_1 and \bar{l}_2 can be inferred, although the corresponding correlation matrix is not given. We will take here two extreme view points, total de-correlation and total anti-correlation. As we will see, the totally anti-correlated case seems to induce smaller errors in the scattering lengths as compared to the totally de-correlated case. We will call these two choices Set **Ia** and Set **Ib** respectively. We will also examine here a partial anti-correlation scenario (Set **Ic**) which models more accurately the limited statistical information provided in Ref. [7]. More details are given in the Appendix.

For the three sets, we take $L_1^r = 0.52 \pm 0.23$, $L_2^r = 0.72 \pm 0.24$ and $L_3^r = -2.69 \pm 0.99$ (main fit in Table 1 of Ref. [7]). For Set **Ia**, we assume that the three statistical distributions are de-correlated, whereas for Set **Ib** [Set **Ic**] assume these are correlated as specified in Eq. (12) [Eq. (13) with $r_{13} = r_{23} = -r_{12} = -0.85$]. By means of a Monte Carlo simulation we generate the statistical distributions for the \bar{l}_1 and \bar{l}_2 parameters, and thus we get

$$\bar{l}_1 = 0.3 \pm 2.1, \quad \bar{l}_2 = 4.77 \pm 0.45, \quad r(\bar{l}_1, \bar{l}_2) = 0, \quad \text{Set **Ia**} \quad (7)$$

$$\bar{l}_1 = 0.3 \pm 1.0, \quad \bar{l}_2 = 4.77 \pm 0.45, \quad r(\bar{l}_1, \bar{l}_2) = -1, \quad \text{Set **Ib**} \quad (8)$$

$$\bar{l}_1 = 0.3 \pm 1.2, \quad \bar{l}_2 = 4.77 \pm 0.45, \quad r(\bar{l}_1, \bar{l}_2) = -0.69, \quad \text{Set **Ic**} \quad (9)$$

being r the linear correlation coefficient. The errors quoted above are clearly low bounds, because they do not account for any systematic effects, in particular those induced by the $\mathcal{O}(p^6)$ -corrections to the relations between the two- and three-flavor low energy constants⁴. Estimates for the systematic errors are not given in Ref. [7] either. Notice that correlations among $L_{1,2,3}^r$ do not affect to the error in \bar{l}_2 (see footnote 4).

⁴Note that in the second entry of Ref. [7] there is an inconsistency. In Eq. (6.24) an error of ± 1.0 for \bar{l}_2 is quoted. That error is not compatible with having $\bar{l}_2 = 192\pi^2 L_2^r - (1 + \ln(m_K^2/\mu^2) + 8 \ln(m^2/\mu^2))/8$, as deduced from Eq. (6.23) of that reference, and an error for L_2^r of $\pm 0.24 \times 10^{-3}$. The previous formula gives

- Set II D -waves: The method of Ref. [4] allows to define another parameter set. If the D -wave $\pi\pi$ scattering lengths for isospin $I = 0$ and $I = 2$, a_{02} and a_{22} respectively, are fixed at the two loop level one gets parameter Set **II**,

$$\bar{l}_1 = -0.8 \pm 4.8, \quad \bar{l}_2 = 4.5 \pm 1.1, \quad r(\bar{l}_1, \bar{l}_2) = -0.75, \quad \text{Set II} \quad (10)$$

To obtain \bar{l}_1 and \bar{l}_2 and their errors, we have propagated the errors in μ , a_{02} , a_{22} , \bar{l}_3 , \bar{l}_4 and the resonance parameters of Eq. (5) in the formula for the D -wave scattering length given in Ref. [4]. This procedure generates a correlation between \bar{l}_1 and \bar{l}_2 which has to be taken into account when calculating errors in quantities depending on the previous parameters. Note that our central values for \bar{l}_1 and \bar{l}_2 numbers are not exactly the ones quoted in Ref. [4] since they take $\mu = 1$ GeV. Furthermore, in Ref. [4] no error estimates are quoted for \bar{l}_1 and \bar{l}_2 . If we had taken $\mu = 1.00 \pm 0.25$ GeV, we would have obtained $\bar{l}_1 = -1.5 \pm 5.8$ and $\bar{l}_2 = 4.5 \pm 1.1$, in agreement with their quoted central value.

- Set III Roy sum rules: Finally, using the method of Roy sum rules, Ref. [8], another parameter set has been obtained. We call it Set **III**,

$$\bar{l}_1 = -0.9 \pm 1.2, \quad \bar{l}_2 = 4.34 \pm 0.25, \quad r(\bar{l}_1, \bar{l}_2) = -0.22, \quad \text{Set III} \quad (11)$$

Both central values and errors quoted above do not agree with those given in Eq. (21) of Ref. [8]. The difference on the central values is due to the fact that in this latter reference all resonance parameters are set to zero at 1 GeV ($r_i(1 \text{ GeV}) = 0 \pm 2 \times 10^{-4}$). To obtain the errors we have propagated the errors in the resonance saturation scale μ , λ_1 , λ_2 (parameters given in Eq. (2) of Ref. [8]), \bar{l}_3 , \bar{l}_4 and the resonance parameters of Eq. (5) when Eqs.(6-7) of Ref. [8] are inverted. This, again generates a correlation between \bar{l}_1 and \bar{l}_2 as in the previous cases⁵. The λ_1 , λ_2 parameters suffer from sizeable systematic uncertainties. Those have been estimated in the second entry of Ref. [3] and turned out to be of comparable magnitude, when not bigger, than the statistical fluctuations, and as a consequence statistical correlations get washed out. This justifies the use of de-correlated distributions for these two parameters both in Ref. [8] and in the present work.

an error for \bar{l}_2 of $\pm 192\pi^2 \times 0.24 \times 10^{-3} = \pm 0.45$ and it is obviously independent of the possible correlations between the $SU(3)$ L_1^r, L_2^r and L_3^r parameters, because it only involves L_2^r . This is in agreement with our results. The problem in Ref. [7] is that the error definition for the $SU(2)$ -parameters, “projections on the relevant variable of the 68% confidence level domain”, is not consistent with that adopted for the $SU(3)$ -parameters. For these latter ones, the standard and traditionally accepted χ^2 -errors are given. The method of the projections would lead to significantly different errors for $L_{1,2,3}^r$, as can be appreciated from the figures shown in Ref. [7]. In anycase, this projection method is not standard. The *standard* and complete procedure consists of giving the full correlation matrix deduced from the χ^2 -fit. We ignore to what extent the errors on derived quantities given in Ref. [7] are affected by this inconsistency.

⁵Errors differ from those quoted in Ref. [8], because of the different treatment considered here (error on the resonance saturation scale, different choice, both for central values and errors, for the resonance parameters but also because there is a numerical mistake in that work, which we correct here. The error analysis of Eqs. (16) and (17) of that work yields $\bar{l}_1 = -0.37 \pm 0.95 \pm 1.27$ and $\bar{l}_2 = 4.17 \pm 0.19 \pm 0.33$, which errors do not agree with those quoted in Eq. (20) of that reference.

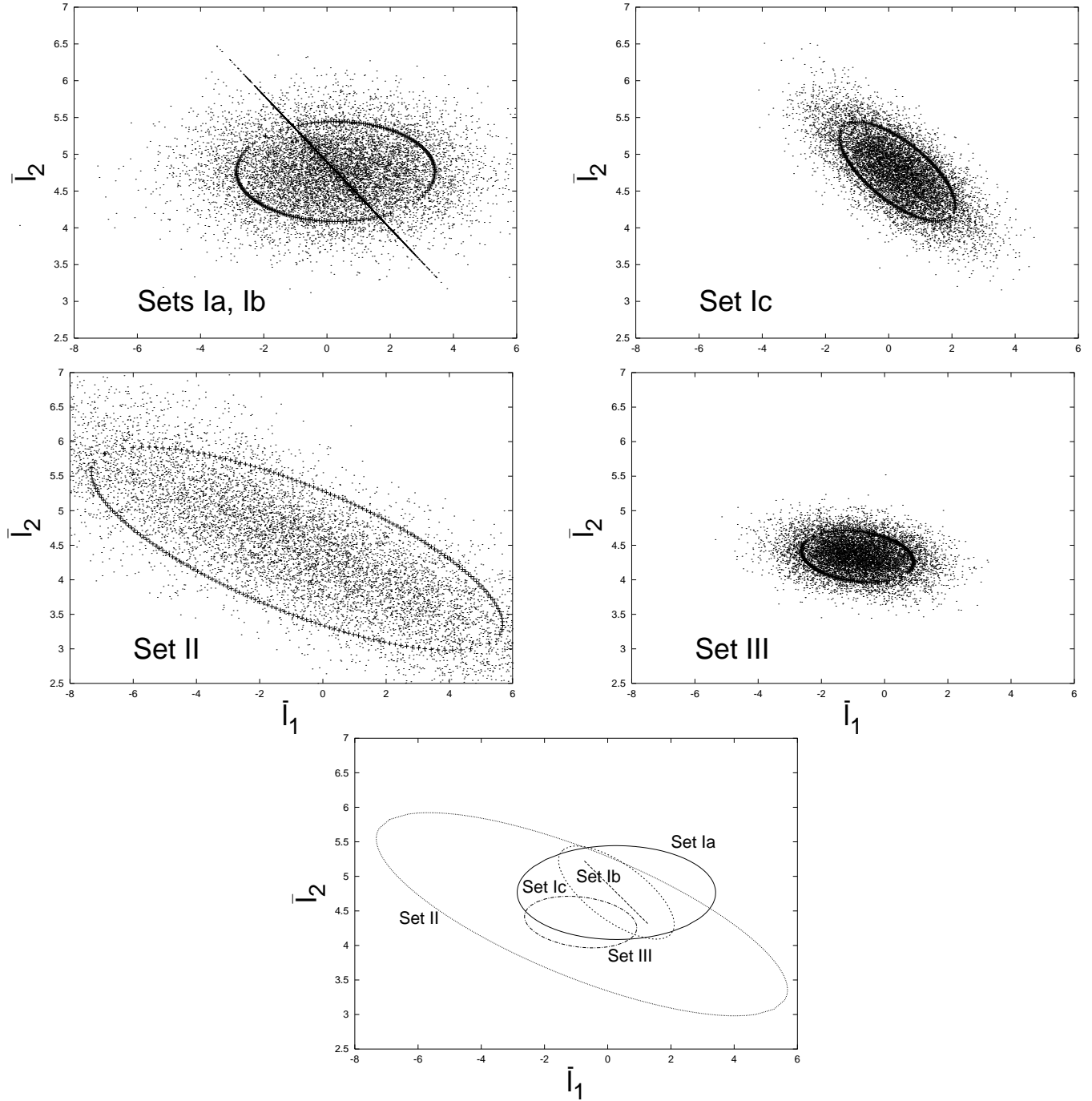


Figure 1: Two dimensional distributions of the low energy parameters \bar{l}_1, \bar{l}_2 generated via Monte Carlo simulation with $N = 10^4$ samples. The contours and the line made of squares for the Set **Ib** represent the 68% confidence limit, i.e., inside them the 68% of the samples are enclosed. Upper left panel, parameter Sets **Ia** and **Ib** corresponding to a two loop K_{l4} form factor analysis [7], with totally de-correlated and totally anti-correlated \bar{l}_1 and \bar{l}_2 parameters respectively. Upper right panel, parameter Set **Ic** corresponding to a two loop K_{l4} form factor analysis [7], with linear correlation $r = -0.69$. Middle left panel, parameter Set **II** corresponding to D -waves in $\pi\pi$ scattering ($r = -0.75$). Middle right panel, parameter Set **III** corresponding to Roy sum rules in $\pi\pi$ scattering ($r = -0.22$). Finally, in the bottom panel the 68% confidence limit contours are superposed.

The determination of the one loop parameters, their corresponding uncertainties as well as their mutual correlation can be best seen by plotting the resulting two-dimensional Monte Carlo generated distributions for all considered sets. This is done in Fig. 1. For all sets we consider a sufficiently large sample, $N = 10^4$, and have checked numerical stability by doubling the sample size. We also plot the 68% confidence limit contours, i.e., inside these boundaries 68% of distribution points are enclosed. For a better comparison we have also superposed (bottom panel) the 68% confidence limit contours to check for statistical compatibility. As could already be anticipated, the main difference between parameter Sets **Ia**, **Ic** and **Ib** (corresponding to K_{l4} decays) is that the errors decrease as the anti-correlation increases. The parameter Set **II** provides the largest uncertainties, and large overlap with the remaining sets is encountered. At the same time, it is clear that there is a trend to discrepancy between parameter Sets **Ic** and **III**, since very little overlap is observed between both distributions. We should also mention that the level of uncertainty regarding both the K_{l4} and the Roy sum rules determinations is comparable, although the latter method seems to be slightly more accurate.

In Ref. [4] the calculation of the two loop contributions to the threshold parameters was undertaken and explicit expressions for scattering lengths and ranges were written, in terms of the low energy constants, $b_{1,2,3,4,5,6}$. In this note we complete their numerical calculation by providing their numbers with the inherited error-bars and also exploring different parameter sets. This is shown in Table 1 (for the b_i parameters) and in Table 2 (for the threshold parameters, a_{IJ}, b_{IJ}).

For all sets, we have generated statistical distributions both in the one loop and the two loop corrections to the scattering lengths and effective ranges by propagating the statistical distributions in the low energy parameters $\bar{l}_{1,2,3,4}$, the resonance parameters $r_{1,2,3,4,5,6}$ and the scale μ discussed above. The uncertainty in the scale and in the resonance contributions, affect to the error of the two loop corrections only. Errors in the two loop contributions are correlated with those in the one loop contributions. Therefore, the error of the sum cannot be obtained by simply adding the errors in quadrature of one and two loop contributions. In any case, for the two loop calculation to be numerically meaningful these one loop uncertainties have to be significantly smaller than the corrections due to the two loop calculation. At the same time the two loop correction has to be significantly smaller than the one loop correction, for a convergent expansion.

As we see in Table 1, all predictions for the b_i parameters are compatible within errors. Moreover, as one might have inferred from the values and distributions of the parameters \bar{l}_1 and \bar{l}_2 the increasing accuracy of the determinations correspond to Sets **II**, **Ia**, **Ic**, **Ib** and **III**, in that order⁶. Besides, we would like to point out that for the Sets **Ia**, **Ib** and **Ic** the statistical fluctuations in the one loop parameters induce errors on the b_i parameters as important, when no more as in the de-correlated case of Set **Ia**, as the scale uncertainty quoted in Ref. [4]. For the parameter Set **II** the central values of the b parameters are slightly different than those in Ref. [4] because the central \bar{l}_1 and \bar{l}_2 values are also different. The errors for the parameter Set **II** are much larger than those found

⁶These two latter parameter sets have similar accuracy.

	Set Ia	Set Ib	Set Ic	Set II	Set III
$10 \cdot b_1$	-0.74 ± 0.21	-0.74 ± 0.05	-0.74 ± 0.16	-0.83 ± 0.48	-0.83 ± 0.15
$10 \cdot b_2$	$+0.71 \pm 0.17$	$+0.71 \pm 0.11$	$+0.71 \pm 0.12$	$+0.78 \pm 0.41$	$+0.78 \pm 0.10$
$10^2 \cdot b_3$	$-0.15 \begin{smallmatrix} +0.30 \\ -0.25 \end{smallmatrix}$	$-0.15 \begin{smallmatrix} +0.19 \\ -0.13 \end{smallmatrix}$	$-0.15 \begin{smallmatrix} +0.21 \\ -0.15 \end{smallmatrix}$	-0.27 ± 0.65	$-0.27 \begin{smallmatrix} +0.19 \\ -0.15 \end{smallmatrix}$
$10^2 \cdot b_4$	$+0.53 \pm 0.07$	$+0.53 \pm 0.06$	$+0.53 \begin{smallmatrix} +0.06 \\ -0.07 \end{smallmatrix}$	$+0.48 \pm 0.13$	$+0.47 \pm 0.03$
$10^3 \cdot b_5$	$+0.22 \begin{smallmatrix} +0.19 \\ -0.24 \end{smallmatrix}$	$+0.22 \begin{smallmatrix} +0.09 \\ -0.14 \end{smallmatrix}$	$+0.22 \begin{smallmatrix} +0.11 \\ -0.16 \end{smallmatrix}$	$+0.09 \begin{smallmatrix} +0.26 \\ -0.35 \end{smallmatrix}$	$+0.07 \begin{smallmatrix} +0.10 \\ -0.13 \end{smallmatrix}$
$10^3 \cdot b_6$	$+0.10 \begin{smallmatrix} +0.03 \\ -0.04 \end{smallmatrix}$	$+0.10 \begin{smallmatrix} +0.03 \\ -0.04 \end{smallmatrix}$	$+0.10 \begin{smallmatrix} +0.03 \\ -0.04 \end{smallmatrix}$	$+0.09 \begin{smallmatrix} +0.02 \\ -0.04 \end{smallmatrix}$	$+0.08 \begin{smallmatrix} +0.03 \\ -0.04 \end{smallmatrix}$

Table 1: Low energy scale independent two loop parameters and their uncertainties due to the error bars in the \bar{l} 's one loop parameters and the uncertainties in both the scale and the resonant part of the two loop contribution. Set **Ia** corresponds to K_{I4} two loop calculation: $\bar{l}_1 = 0.3 \pm 2.1$, $\bar{l}_2 = 4.77 \pm 0.45$, $\bar{l}_3 = 2.9 \pm 2.4$, $\bar{l}_4 = 4.4 \pm 0.3$, with \bar{l}_1 and \bar{l}_2 uncorrelated. Set **Ib** corresponds to K_{I4} two loop calculation: $\bar{l}_1 = 0.3 \pm 1.0$, $\bar{l}_2 = 4.77 \pm 0.46$, $\bar{l}_3 = 2.9 \pm 2.4$, $\bar{l}_4 = 4.4 \pm 0.3$, with \bar{l}_1 and \bar{l}_2 totally anti-correlated. Set **Ic** corresponds to K_{I4} two loop calculation: $\bar{l}_1 = 0.3 \pm 1.2$, $\bar{l}_2 = 4.77 \pm 0.44$, $\bar{l}_3 = 2.9 \pm 2.4$, $\bar{l}_4 = 4.4 \pm 0.3$, with a correlation coefficient $r(\bar{l}_1, \bar{l}_2) = -0.69$. Set **II** corresponds to a two loop D -wave $\pi\pi$ scattering lengths calculation (see main text): $\bar{l}_1 = -0.8 \pm 4.8$, $\bar{l}_2 = 4.5 \pm 1.1$, $\bar{l}_3 = 2.9 \pm 2.4$, $\bar{l}_4 = 4.4 \pm 0.3$, with a correlation coefficient $r(\bar{l}_1, \bar{l}_2) = -0.75$. Set **III** corresponds to a Roy equation sum rule analysis of $\pi\pi$ scattering at two loops: $\bar{l}_1 = -0.9 \pm 1.2$, $\bar{l}_2 = 4.34 \pm 0.25$, $\bar{l}_3 = 2.9 \pm 2.4$, $\bar{l}_4 = 4.4 \pm 0.3$, with a correlation coefficient $r(\bar{l}_1, \bar{l}_2) = -0.22$. The scale is $\mu = 750 \pm 250$ MeV under the resonance saturation hypothesis, for which a 100% error is assumed.

for parameter Sets **I** and **III**. The size of the errors for Set **II** is comparable with the ones for b_3, b_4, b_5, b_6 found in Ref. [3]⁷.

As we see in Table 2, results for a_{20} , b_{20} and a_{22} with parameter Set **Ia** are such that the one loop errors are larger than the central values of the two loop contribution. The inclusion of statistical correlations, Sets **Ib** and **Ic**, increases the accuracy of the predictions, though it still happens that errors on the one loop contribution make irrelevant the two loop one. The situation worsens dramatically, for parameter Set **II** where we see that in most considered cases predictive power is lost beyond one loop, with the exception of the S -wave scattering lengths a_{00} and a_{02} . Set **III**, obtained from the Roy-sum rule analysis of Ref. [8], turns out to be as predictive as Sets **Ib** and **Ic**, obtained from the two-loop improved K_{I4} analysis of Ref. [7]. In any case is also true that for a_{20} , b_{20} and a_{22} within Set **III** the one loop uncertainties are larger than the two loop contribution. The results of the table are compatible, for all parameter sets with the experimental analysis of $\pi\pi$ scattering data [11] but produce in general much better errors despite of the

⁷ Regarding the parameter Set **III**, one should say that if one would calculate the b 's from Eq. (48) of Ref. [3], with the numerical input of Eq.(2) of Ref. [8], one would have obtained

$$\begin{aligned}
10 b_1 &= -0.68 \pm 0.09 \quad , \quad 10 b_2 = +0.64 \pm 0.09 \quad , \quad 10^2 b_3 = -0.35 \pm 0.24 \\
10^2 b_4 &= +0.47 \pm 0.03 \quad , \quad 10^3 b_5 = 0.13 \pm 0.06 \quad , \quad 10^3 b_6 = 0.10 \pm 0.01
\end{aligned}$$

$b_{3,4,5,6}$ are in agreement with Eq.(49) of Ref. [3]. We see that this is *not the same* as evaluating the b 's in the spirit of resonance saturation (Eqs. D.1,D.2 and D.3 of second entry in Ref. [4]) from previously computed values of \bar{l}_1 and \bar{l}_2 , through their relation to λ_1 and λ_2 parameters, as suggested in Ref [8] and adopted in this work. In any case, the values given above and those corresponding to Set **III** in Table 1 are compatible within statistical fluctuations.

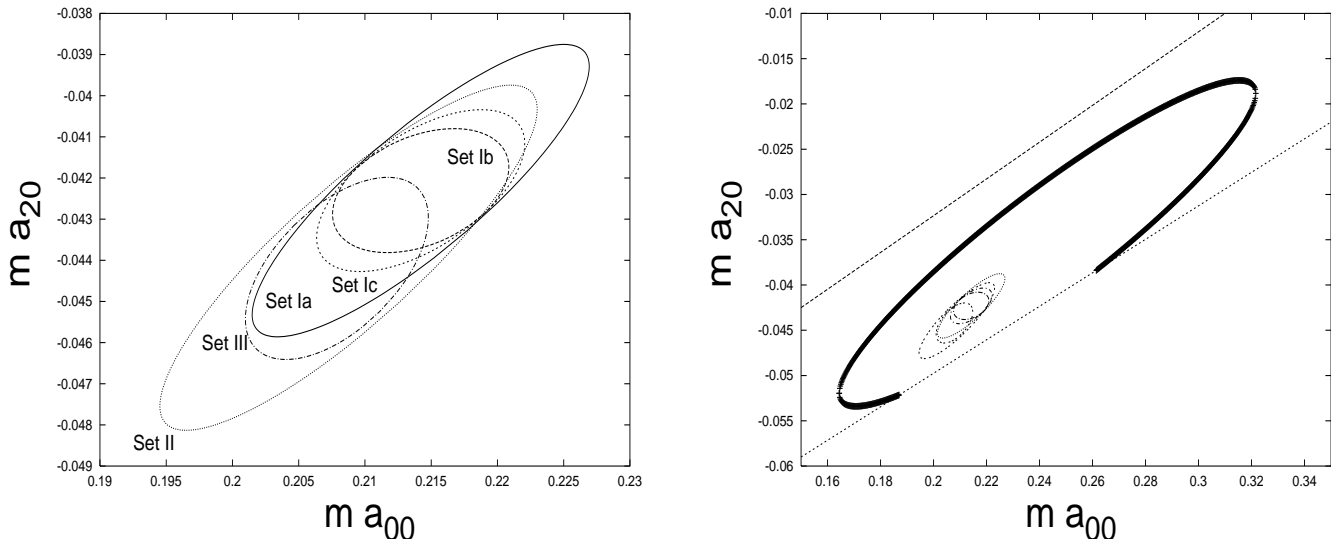


Figure 2: Left panel: contours corresponding to a 68% confidence limits for the two dimensional distributions of the S -wave scattering lengths a_{00} and a_{20} . Parameter sets are defined in the main text. The linear correlation coefficients, $r(a_{00}, a_{20})$, for Sets **Ia**, **Ib**, **Ic**, **II** and **III** are 0.85, 0.38, 0.58, 0.86 and 0.56 respectively. Right panel: The same as before with the additional inclusion of the universal bands and the 70% confidence limit contour (crosses) deduced from a numerical analysis of the Roy equations in Ref. [12].

problems discussed above on the relevance of the two loop contributions. The early work on K_{l4} form factors using a dispersive one loop calculation for three flavors [6], predicted a isoscalar D -wave three standard deviations above the value extracted from the experiment [8], the upgrade of this calculation to two loop accuracy in Ref. [7] have contributed to considerably improve such a discrepancy.

In Fig. 2 we exhibit the Monte Carlo propagated two dimensional distributions of the S -wave scattering lengths, for isospin channels $I = 0$ and $I = 2$, for all parameter sets and in terms of the 68% confidence limit contours. It is interesting to compare this distributions with the preliminary results of Ref. [12], based on the detailed numerical analysis of Roy equations in $\pi\pi$ scattering. Such an analysis implies the existence of a universal band, outside which no solution to Roy equations exists, as a mathematical consequence of analyticity, crossing, unitarity and isospin conservation. In addition to this, the physical requirement of describing the available experimental data, provides a 70% confidence limit contour [12]. As can be seen in the figure, with all parameter sets considered in this work, the two-loop ChPT analysis are not only compatible with the 70% confidence limit contour of Ref. [12] but have significant smaller statistical fluctuations.

In summary, effective field theories like Chiral Perturbation Theory have predictive power, but it is not unlimited because of three reasons: 1) truncation of the expansion, 2) proliferation of undetermined constants at any order and 3) experimental uncertainties. Thus, experimental data prove crucial to determine the, increasing with the order, unknown constants and their errors, which propagate in a correlated way to higher orders in the expansion possibly undermining the “convergence” of the expansion. This situation also appears in fundamental theories like QED or QCD, but it is in fact worse in ChPT

because the number of unknown parameters in these theories does not increase with the order of the expansion. This why error analysis is so important. We have exemplified our points in the calculation of the threshold parameters for $\pi\pi$ scattering up to two loops. The general picture provided by ChPT is rather satisfactory, in the sense that the accuracy of the predictions is much bigger than the available data ⁸. Nevertheless we find cases where, within the present experimental accuracy, the errors in the one loop contribution are larger than the central values of the two loop contribution. Conclusions regarding the loss of predictive power, can only be reinforced if systematic uncertainties induced by the $\mathcal{O}(p^6)$ -corrections to the relations between the two- and three-flavor low energy constants are included in the results of Ref. [7]. The effect of these systematic corrections will be twofold: a general increase of the error fluctuations and a decrease of the correlations. Both effects will contribute to increase the errors on all derived threshold and low energy parameters presented in this work and deduced from K_{l4} -decays.

ChPT defines a whole family of effective theories, but obviously the most interesting choice exactly corresponds to low energy QCD. In QCD with two flavors and neglecting isospin breaking there appear only two independent parameters, Λ_{QCD} and the quark mass m_Q . The quantities f_π , m , the \bar{l} 's, the b 's and higher order parameters must be functions of them. This dependence introduces correlations among all low energy parameters which, if taken into account, would influence the present error analysis, and presumably might yield to more moderate errors in terms of Λ_{QCD} and m_Q and their errors. Assuming more statistically independent parameters than QCD suggests is, so far, another manifestation of the inability to undertake a quantitative and microscopic derivation of ChPT as an effective low energy theory of QCD on the one hand, but on the other hand reassures ChPT as a convenient tool to deal with non-perturbative phenomena in strong interactions.

Appendix: Modeling statistical correlations in K_{l4} decays.

In Ref. [7] is stated that L_3^r is strongly anti-correlated with L_1^r and L_2^r , though central values and errors for these parameters are given ($L_1^r = 0.52 \pm 0.23$, $L_2^r = 0.72 \pm 0.24$ and $L_3^r = -2.69 \pm 0.99$, main fit in Table 1 of Ref. [7]), however the correlation matrix is not provided. Besides, new parameters $X_1 = L_2^r - 2L_1^r - L_3^r$, $X_2 = L_2^r$ and $X_3 = (L_2^r - 2L_1^r)/L_3^r$ are introduced, it is quoted the value $X_3 = 0.12_{-11}^{+8}$ and it is also specified that X_3 is little correlated with X_1 and X_2 . We will show below how statistical modeling and the fragmentary information given in Ref. [7] can be used to reconstruct some relevant information on the correlations between \bar{l}_1 and \bar{l}_2 . To be most objective, we also explore two other extreme cases which correspond to total de-correlation and total anti-correlation. This said, we explore three different scenarios in this work:

- A total de-correlated picture ($r(L_1^r, L_3^r) = r(L_2^r, L_3^r) = r(L_1^r, L_2^r) = 0$, being r the linear correlation coefficient). It leads to $X_3 = 0.12_{-20}^{+23}$, $r(X_1, X_3) = -0.14$ and $r(X_2, X_3) = -0.09$.

⁸Right plot of Fig.2 constitutes a clear example.

- A total anti-correlation scenario ($r(L_1^r, L_3^r) = r(L_2^r, L_3^r) = -1$) which can be implemented by simply assuming a linear relation

$$(L_1^r - \langle L_1^r \rangle) / \sigma_1 = -(L_3^r - \langle L_3^r \rangle) / \sigma_3 = (L_2^r - \langle L_2^r \rangle) / \sigma_2, \quad (12)$$

That is the necessary and sufficient condition for total anti-correlation. Besides, one also has $r(L_1^r, L_2^r) = +1$. This simple model leads to $X_3 = 0.12_{-6}^{+3}$ and $r(X_1, X_3) = r(X_2, X_3) = 0.06$. Thus, the total anti-correlation scenario, though simple, provides an acceptable description, and in any case much more precise than when correlations between L_1^r , L_2^r and L_3^r are neglected, of the findings of Ref. [7].

- A partial anti-correlation scenario which can be implemented if one assumes that L_1^r , L_2^r and L_3^r are gaussian distributed⁹ according to

$$P(L_1^r, L_2^r, L_3^r) = \frac{(\det[C])^{\frac{1}{2}}}{(2\pi)^{\frac{3}{2}}\sigma_1\sigma_2\sigma_3} \times e^{-\frac{1}{2}(L^T \cdot C \cdot L)} \quad (13)$$

where P is the joint density probability distribution of the three random variables, σ_i is the L_i^r error, L is a column matrix with elements $(L_i^r - \mu_i) / \sigma_i$, being μ_i the central value of the variable L_i^r , and finally the inverse of the symmetric matrix C is given by

$$C^{-1} = \begin{pmatrix} 1 & r_{12} & r_{13} \\ r_{12} & 1 & r_{23} \\ r_{13} & r_{23} & 1 \end{pmatrix} \quad (14)$$

and therefore it is determined by the linear correlation coefficients $r_{ij} = r(L_i^r, L_j^r)$. Our aim is to improve the total anti-correlated scenario presented in the item above but keeping the model still as simple as possible. Thus we have explored situations where $r_{13} = r_{23} = -r_{12} = r$, for which one can analytically diagonalize the C -matrix in Eq. (13). We find for $r = -0.85$, still a strong anti-correlation, $X_3 = 0.12_{-10}^{+8}$, $r(X_1, X_3) = -0.02$ and $r(X_2, X_3) = -0.01$. Thus, this model for the correlations reproduces pretty much the results given in the K_{14} analysis of Ref. [7]. This can also be appreciated in Fig. 3 where we show the resulting two dimensional distributions of the $SU(3)$ -low energy parameters L_1^r, L_3^r and L_2^r, L_3^r . These distributions compare reasonably well with those given in Fig.2 of the first entry of Ref. [7]. This makes us more confident on the validity of the simple statistical model used here.

The correlations between L_1^r , L_2^r and L_3^r are transported through the equations relating $SU(2)$ and $SU(3)$ low energy constants. For the total anti-correlated scenario leads to a total anti-correlation for \bar{l}_1 and \bar{l}_2 , it is to say $(\bar{l}_1 - \langle \bar{l}_1 \rangle) / \sigma_1 = -(\bar{l}_2 - \langle \bar{l}_2 \rangle) / \sigma_2$ and therefore $r(\bar{l}_1, \bar{l}_2) = -1$. For the total de-correlated case the correlation coefficient $r(\bar{l}_1, \bar{l}_2)$ is zero, whereas for the partial correlation scenario we get $r(\bar{l}_1, \bar{l}_2) = -0.69$.

⁹This is totally justified, because in Ref. [7] these variables have been determined from a χ^2 -fit.

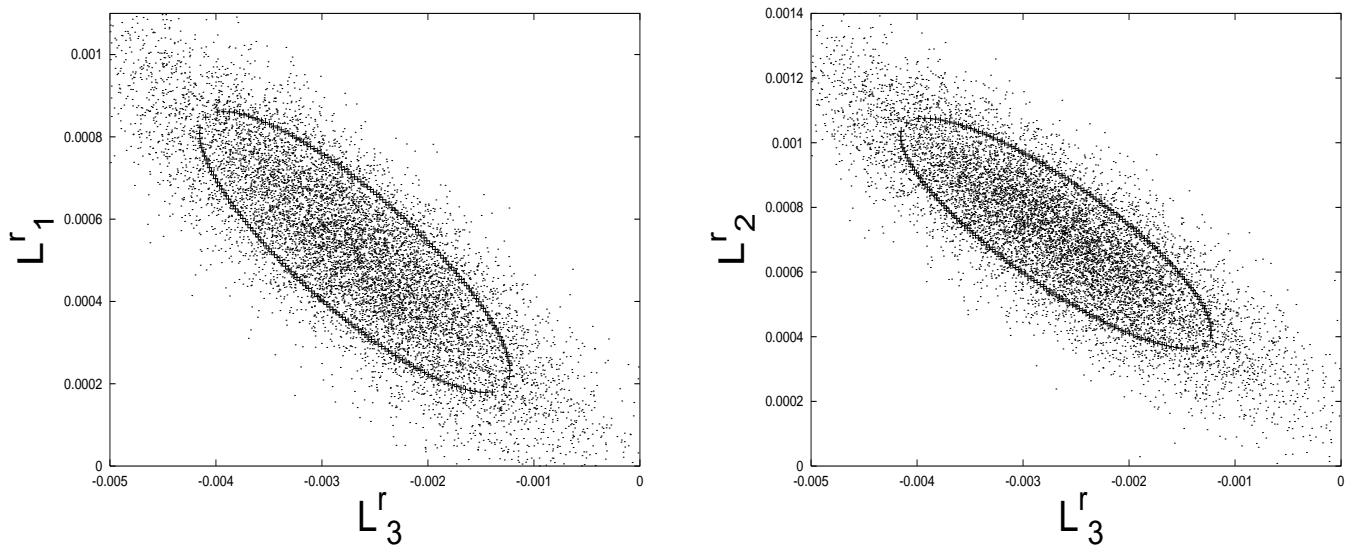


Figure 3: Two dimensional distributions of the $SU(3)$ -low energy parameters L_1^r, L_3^r and L_2^r, L_3^r . generated via Monte Carlo simulation with $N = 10^4$ samples. The contours represent the 68% confidence limit, i.e., inside them the 68% of the samples are enclosed.

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$a_{IJ}; b_{IJ}$	Set	(tree)	+(1loop)	+(2loop)	total	experiment
$a_{00}m$	Ia	0.156	0.043 ± 0.005	0.015 ± 0.004	0.214 ± 0.008	0.26 ± 0.05
	Ib	“	“ ± 0.002	“ ± 0.003	“ ± 0.004	“
	Ic	“	“ ± 0.003	“ ± 0.003	“ ± 0.005	“
	II	“	0.040 ± 0.007	0.013 ± 0.004	0.209 ± 0.010	“
	III	“	0.039 ± 0.003	0.012 ± 0.002	0.208 ± 0.005	“
$b_{00}m^3$	Ia	0.179	0.070 ± 0.015	0.025 ± 0.011	0.273 ± 0.024	0.25 ± 0.03
	Ib	“	“ ± 0.003	“ $\begin{smallmatrix} + 0.007 \\ - 0.008 \end{smallmatrix}$	“ $\begin{smallmatrix} + 0.008 \\ - 0.009 \end{smallmatrix}$	“
	Ic	“	“ ± 0.006	“ ± 0.008	“ $\begin{smallmatrix} + 0.011 \\ - 0.012 \end{smallmatrix}$	“
	II	“	0.059 ± 0.024	0.019 ± 0.013	$0.257 \begin{smallmatrix} + 0.032 \\ - 0.036 \end{smallmatrix}$	“
	III	“	0.058 ± 0.008	0.018 ± 0.006	0.254 ± 0.010	“
$10 \cdot a_{11}m^3$	Ia	0.297	0.055 ± 0.012	$0.021 \begin{smallmatrix} + 0.003 \\ - 0.004 \end{smallmatrix}$	$0.373 \begin{smallmatrix} + 0.011 \\ - 0.012 \end{smallmatrix}$	0.38 ± 0.02
	Ib	“	“ ± 0.009	“ $\begin{smallmatrix} + 0.002 \\ - 0.003 \end{smallmatrix}$	“ ± 0.011	“
	Ic	“	“ ± 0.009	“ $\begin{smallmatrix} + 0.002 \\ - 0.003 \end{smallmatrix}$	“ ± 0.011	“
	II	“	0.059 ± 0.033	$0.018 \begin{smallmatrix} + 0.003 \\ - 0.004 \end{smallmatrix}$	$0.375 \begin{smallmatrix} + 0.035 \\ - 0.031 \end{smallmatrix}$	“
	III	“	0.059 ± 0.007	0.018 ± 0.002	0.374 ± 0.007	“
$10 \cdot b_{11}m^5$	Ia	0	0.030 ± 0.012	$0.025 \begin{smallmatrix} + 0.006 \\ - 0.008 \end{smallmatrix}$	$0.055 \begin{smallmatrix} + 0.008 \\ - 0.011 \end{smallmatrix}$	—
	Ib	“	“ $\begin{smallmatrix} + 0.008 \\ - 0.009 \end{smallmatrix}$	“ $\begin{smallmatrix} + 0.002 \\ - 0.004 \end{smallmatrix}$	“ $\begin{smallmatrix} + 0.009 \\ - 0.011 \end{smallmatrix}$	—
	Ic	“	“ ± 0.009	“ $\begin{smallmatrix} + 0.003 \\ - 0.005 \end{smallmatrix}$	“ $\begin{smallmatrix} + 0.009 \\ - 0.011 \end{smallmatrix}$	—
	II	“	0.034 ± 0.033	$0.020 \begin{smallmatrix} + 0.005 \\ - 0.009 \end{smallmatrix}$	0.054 ± 0.029	—
	III	“	0.034 ± 0.007	$0.019 \begin{smallmatrix} + 0.003 \\ - 0.005 \end{smallmatrix}$	$0.053 \begin{smallmatrix} + 0.005 \\ - 0.006 \end{smallmatrix}$	—
$10 \cdot a_{20}m$	Ia	-0.446	0.021 ± 0.021	0.002 ± 0.004	-0.423 ± 0.024	-0.28 ± 0.12
	Ib	“	“ ± 0.008	“ ± 0.004	“ ± 0.010	“
	Ic	“	“ ± 0.011	“ ± 0.004	“ ± 0.013	“
	II	“	0.007 ± 0.028	0.000 ± 0.004	$-0.439 \begin{smallmatrix} + 0.028 \\ - 0.030 \end{smallmatrix}$	“
	III	“	0.004 ± 0.013	$-0.001 \begin{smallmatrix} + 0.003 \\ - 0.004 \end{smallmatrix}$	-0.442 ± 0.015	“
$10 \cdot b_{20}m^3$	Ia	-0.892	0.129 ± 0.043	$0.003 \begin{smallmatrix} + 0.013 \\ - 0.011 \end{smallmatrix}$	-0.760 ± 0.043	-0.82 ± 0.08
	Ib	“	“ ± 0.010	“ $\begin{smallmatrix} + 0.012 \\ - 0.010 \end{smallmatrix}$	“ $\begin{smallmatrix} + 0.016 \\ - 0.014 \end{smallmatrix}$	“
	Ic	“	“ $\begin{smallmatrix} + 0.018 \\ - 0.019 \end{smallmatrix}$	“ $\begin{smallmatrix} + 0.013 \\ - 0.011 \end{smallmatrix}$	“ $\begin{smallmatrix} + 0.023 \\ - 0.021 \end{smallmatrix}$	“
	II	“	0.095 ± 0.045	$0.003 \begin{smallmatrix} + 0.016 \\ - 0.013 \end{smallmatrix}$	-0.794 ± 0.039	“
	III	“	0.088 ± 0.022	0.003 ± 0.011	-0.801 ± 0.021	“
$10^2 \cdot a_{02}m^5$	Ia	0	0.143 ± 0.031	$0.058 \begin{smallmatrix} + 0.016 \\ - 0.019 \end{smallmatrix}$	0.202 ± 0.042	0.17 ± 0.03
	Ib	“	“ ± 0.009	“ $\begin{smallmatrix} + 0.016 \\ - 0.019 \end{smallmatrix}$	“ $\begin{smallmatrix} + 0.021 \\ - 0.025 \end{smallmatrix}$	“
	Ic	“	“ ± 0.015	“ $\begin{smallmatrix} + 0.016 \\ - 0.019 \end{smallmatrix}$	“ $\begin{smallmatrix} + 0.024 \\ - 0.028 \end{smallmatrix}$	“
	II	“	0.117 ± 0.028	0.053 ± 0.022	0.170 ± 0.030	“
	III	“	0.111 ± 0.016	$0.051 \begin{smallmatrix} + 0.013 \\ - 0.016 \end{smallmatrix}$	0.163 ± 0.020	“
$10^3 \cdot a_{22}m^5$	Ia	0	0.278 ± 0.241	$-0.035 \begin{smallmatrix} + 0.09 \\ - 0.07 \end{smallmatrix}$	0.24 ± 0.17	0.13 ± 0.30
	Ib	“	“ ± 0.062	“ $\begin{smallmatrix} + 0.057 \\ - 0.039 \end{smallmatrix}$	“ $\begin{smallmatrix} + 0.07 \\ - 0.05 \end{smallmatrix}$	“
	Ic	“	“ ± 0.108	“ $\begin{smallmatrix} + 0.064 \\ - 0.044 \end{smallmatrix}$	“ $\begin{smallmatrix} + 0.10 \\ - 0.08 \end{smallmatrix}$	“
	II	“	0.119 ± 0.460	$0.011 \begin{smallmatrix} + 0.172 \\ - 0.124 \end{smallmatrix}$	0.13 ± 0.30	“
	III	“	0.102 ± 0.130	$0.015 \begin{smallmatrix} + 0.072 \\ - 0.061 \end{smallmatrix}$	0.12 ± 0.11	“

Table 2: Threshold $\pi\pi$ scattering parameters and their uncertainties in units of m due to the error bars in the \bar{l}^s one loop parameters and the uncertainties in both the scale and the resonant part of the two loop contribution. Sets **Ia**, **Ib**, **Ic**, **II** and **III** are defined in table 1 and main text. Errors are *not* added in quadrature due to statistical correlations. Experimental values are from Ref. [11].