

# Familial confounding of the association between maternal smoking during pregnancy and ADHD in offspring

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**Background:** Maternal Smoking During Pregnancy (SDP) has consistently been associated with increased risk of attention-deficit/hyperactivity disorder (ADHD) in offspring, but recent studies indicate that this association might be due to unmeasured familial confounding. **Methods:** A total of 813,030 individuals born in Sweden between 1992 and 2000 were included in this nationwide population-based cohort study. Data on maternal SDP and ADHD diagnosis were obtained from national registers and patients were followed up from the age of 3 to the end of 2009. Hazard Ratios (HRs) were estimated using stratified Cox regression models. Cousin and sibling data were used to control for unmeasured familial confounding. **Results:** At the population level maternal SDP predicted ADHD in offspring ( $HR_{\text{ModerateSDP}} = 1.89$ ;  $HR_{\text{HighSDP}} = 2.50$ ). This estimate gradually attenuated toward the null when adjusting for measured confounders ( $HR_{\text{ModerateSDP}} = 1.62$ ;  $HR_{\text{HighSDP}} = 2.04$ ), unmeasured confounders shared within the extended family (i.e., cousin comparison) ( $HR_{\text{ModerateSDP}} = 1.45$ ;  $HR_{\text{HighSDP}} = 1.69$ ), and unmeasured confounders within the nuclear family (i.e., sibling comparison) ( $HR_{\text{ModerateSDP}} = 0.88$ ;  $HR_{\text{HighSDP}} = 0.84$ ). **Conclusions:** Our results suggest that the association between maternal SDP and offspring ADHD are due to unmeasured familial confounding. **Keywords:** Maternal smoking during pregnancy, attention-deficit/hyperactivity disorder, confounding, sibling comparisons.

## Introduction

Attention-Deficit/Hyperactivity Disorder (ADHD), characterized by impairing symptoms of hyperactivity, impulsivity and inattention, is a highly prevalent neuropsychiatric disorder that persists into adulthood in a sizable number of affected children of both genders (Biederman, 2005). In addition to a strong genetic predisposition to ADHD, environmental factors account for an estimated 10%–40% of the variance in liability to the disorder (Banerjee, Middleton, & Faraone, 2007). Maternal lifestyle factors during pregnancy, such as stress, alcohol use, and cigarette smoking may represent environmental risk factors for ADHD in offspring (Thapar, Cooper, Eyre, & Langley, 2012), but the mechanisms through which such risks influence ADHD is poorly understood.

Studies have shown that prenatal exposure to maternal Smoking During Pregnancy (SDP) is associated with low birthweight and preterm birth (Cnattingius, 2004) and also suggest plausible biological mechanisms through which SDP may influence brain development, including deleterious effects on neurotransmission, neuronal differentiation, and migration (Dwyer, McQuown, & Leslie, 2009; Slotkin, Tate, Cousins, & Seidler, 2006). A robust association between SDP and ADHD in offspring

has been observed in previous epidemiological studies, even after adjustment for measured confounders (Banerjee et al., 2007; Langley, Rice, van den Bree, & Thapar, 2005; Linnet et al., 2003; Motlagh et al., 2010; Rodriguez & Bohlin, 2005) suggesting possible causality. Other researchers express great skepticism, however, because of the inability of these studies to rule out unmeasured familial factors (Knopik, 2009; Thapar & Rutter, 2009). For instance, the heritability of both SDP (Agrawal et al., 2008; D'Onofrio et al., 2003; Ellingson, Rickert, Lichtenstein, Langstrom, & D'Onofrio, 2012) and ADHD (Larsson, Anckarsater, Rastam, Chang, & Lichtenstein, 2012), as well as the recent reports of a genetic overlap between ADHD and smoking (August et al., 2006; Chang, Lichtenstein, & Larsson, 2012; Monuteaux et al., 2008), indicate that maternal SDP may reflect a genetic predisposition rather than a causal risk factor for offspring ADHD. As randomized controlled trials are impossible, quasi-experimental designs (e.g., sibling and cousin comparison analyses) are needed to rule out unmeasured familial factors (Rutter, 2009).

Available evidence from studies using different quasi-experimental designs including one study of SDP in both mothers and fathers (Langley, Heron, Smith, & Thapar, 2012), one in vitro cross-fostering study (Thapar et al., 2009), and three sibling comparison studies (D'Onofrio et al., 2008; Lindblad & Hjern, 2010; Obel et al., 2011) suggest that the

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dose-dependent relationship between SDP and offspring ADHD is due to unmeasured familial factors (i.e., genetic and/or household-level factors). However, these previous quasi-experimental studies of ADHD have noteworthy limitations. First, studies comparing the offspring ADHD risk between maternal and paternal SDP have produced mixed results and are limited by small sample sizes (Langley et al., 2012; Nomura, Marks, & Halperin, 2010). Second, using children who remain unrelated to their mothers through assisted conception to study SDP is a novel approach for disentangling genetic confounds from prenatal risks, but the findings for SDP and ADHD need to be replicated, as the results were based on parent-rated ADHD symptoms during childhood from a small study of children conceived through assisted reproductive technologies (Thapar et al., 2009). Third, previous sibling studies have been limited by imprecise risk estimates (because of small sample sizes) and by inadequate proxy measures, including use of ADHD medication during only one calendar year (Lindblad & Hjern, 2010) and maternal ratings of ADHD-related problems (D'Onofrio et al., 2008). In addition, previous sibling studies were not able to identify and exclude half-siblings which mean that the effect of familial confounding may have been underestimated (D'Onofrio et al., 2008; Obel et al., 2011). Clearly, additional quasi-experimental research is needed to address these limitations.

Using a large population-based cohort of children, we tested the hypothesis that maternal SDP is a causal risk factor for offspring ADHD. To explicitly address concerns about the generalizability of the findings from previous sibling comparison studies (Talati & Weissman, 2010) and limitations inherent in the design (e.g., women who vary in their smoking status are different from those that smoke across pregnancies) (D'Onofrio, Lahey, Turkheimer, & Lichtenstein, 2013) the current study also included cousin comparisons. We explored the impact of confounding factors by (a) controlling for measured covariates, (b) comparing full cousins (offspring of adult full siblings) within extended families to control for all unmeasured factors that make cousins similar, and (c) comparing full siblings within nuclear families to control for all unmeasured factors that make siblings similar.

## Method

### Sample

**Data sources.** We utilized data from a record linkage of eight population-based registries in Sweden; personal identification numbers enabled accurate linkage.

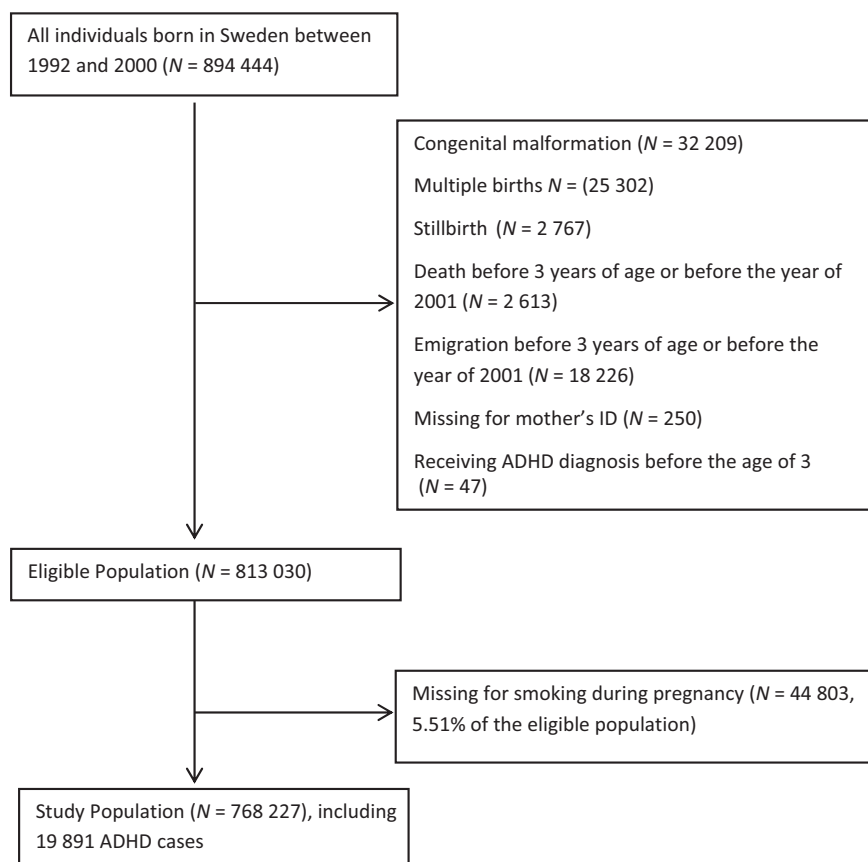
The population-based Medical Birth Register (MBR) provides information on SDP in Sweden since 1983 through self-reported prospective information

at the registration to antenatal care (in pregnancy week 8–12). Swedish maternal health care has a national homogenous structure and reaches almost 99% of all Swedish women before pregnancy week 15 (Lindmark & Cnattingius, 1991). The Patient Register (PR) provides data on psychiatric inpatient care since 1987 (ICD-9 to ICD-10) and outpatient care (ICD-10) since 2001. The Swedish Prescribed Drug register (PDR) (Wettermark et al., 2007) contain information on drug identity (Anatomical Therapeutic Chemical [ATC-codes]) and dates of all registered prescriptions to the entire population in Sweden since July 2005. The Multigeneration Register (MGR) contain information on the identity of the parents of all residents born in Sweden since 1932. The Cause of Death Register (CDR) provides information on dates of all registered deaths since 1958. The Migration Register (MR) includes information on dates of all registered migrations into or out of Sweden since 1969. The LISA Register contains information on education, employment, and income of individuals in each household in Sweden at the end of 2009. The Stockholm Children and Adolescents Psychiatric Care Register (Pastill) cover information on psychiatric diagnoses based on both ICD-10 and Diagnostic and Statistical Manual of Mental Disorders (DSM-IV) for all children and adolescents living in Stockholm County since 2001. The study was approved by the research ethics committee at Karolinska Institute, Stockholm, Sweden Protocol nr 2009/5:9.

**Study population.** After identification of all individuals born in Sweden between 1992 and 2000 ( $N = 894,444$ ) and exclusion of individuals with serious congenital malformations ( $N = 32,209$ ), multiple births ( $N = 25,302$ ), still born before or during delivery ( $N = 2,767$ ), dead before 3 years of age or before 2001 ( $N = 2,613$ ), emigrated before 3 years of age or before 2001 ( $N = 18,226$ ), missing data on mother's identification number ( $n = 250$ ) or who had received an ADHD diagnosis before 3 years of age ( $n = 47$ ), an eligible sample of 813,030 individuals (90.90% of the targeted population) was identified from MBR (Figure 1). After exclusion due to missing values on SDP ( $N = 44,803$ ) the study population included 768,227 individuals covering 365,442 full siblings nested within 172,701 families and 155,852 cousins nested within 52,183 families. All the individuals were followed up from the age of 3 until diagnosis of ADHD, death, emigration, or 31 December 2009, whichever occurred first.

### Measures

**Offspring Outcomes.** We identified 13,996 patients with an in- or outpatient diagnosis of Hyperkinetic Disorder (HKD) (F90 in ICD-10) between January 2001 and December 2009 from PR, and 3,352 patients with diagnoses of HKD (F90 in ICD-10) or ADHD (DSM-IV:314) from Pastill. A total of 16,216



**Figure 1** Sample selection

patients treated with stimulant or nonstimulant medication for ADHD (methylphenidate [N06BA04]; atomoxetine [N06BA09]; amphetamine [N06BA01]; dexamphetamine [N06BA02]) at any time between July 2005 and December 2009 were identified via the PDR. Altogether 19,891 unique ADHD cases were identified after adjustment of overlaps between the three registers. According to Swedish national clinical guidelines, issued by The National Board of Health and Welfare, ADHD or HKD should be diagnosed by a specialist psychiatrist after a clinical somatic and psychiatric evaluation including childhood history, somatic evaluation, and cognitive assessments by a child psychologist. ADHD and HKD diagnoses in PR as well as in Pastill are based on DSM-IV and/or ICD-10 criteria. We used data from 19,150 twins (born between 1992 and 2001) with psychiatric symptom information from the Swedish Twin Registry to check the validity of the register-based ADHD diagnosis. ADHD symptoms were assessed using a well-validated measure of 96 specific child psychiatric symptoms (Hansson et al., 2005; Larson et al., 2010). About 70% of the twins with a national register-based ADHD diagnosis recorded in the National Patient Register were also rated as screen positive by parents (Larsson et al., 2013). Because it is well established that discrepancies often exist among different raters (e.g., parents, children, physicians, and psychologists) (De Los Reyes & Kazdin, 2005), the high overlap between

different informants suggest that the national register-based diagnoses of ADHD are valid.

**Exposure.** SDP was measured on a 3-point scale (No SDP = 0, moderate SDP = 1–9, or high SDP =  $\geq 10$  cigarettes per day). The measure has been validated using numerous outcome measures, and has shown to be highly correlated with serum cotinine levels (Lindqvist, Lendahls, Tollbom, Aberg, & Hakansson, 2002). Previous studies have shown that the validity of self-reported smoking has not changed over time and also acceptable validity of self-reported smoking habits when compared with biochemical markers (measured levels of cotinine in blood) (Haziandreu et al., 1989).

**Covariates.** Based on previous research (D'Onofrio, Singh, Iliadou, Lambe, Hultman, Grann et al., 2010; D'Onofrio, Singh, Iliadou, Lambe, Hultman, Neiderhiser et al., 2010; Lundberg et al., 2010), measured covariates included sex, birth year (1992–1994, 1995–1997, and 1998–2000), mother's parity (1st, 2nd, 3rd, or  $\geq 4$ th), maternal age at childbirth ( $\leq 19$ , 20–24, 25–29, 30–34, or  $\geq 35$  years), cohabitation with offspring's father (yes or no), maternal highest education ( $\leq 9$  years, 10–12 years or graduation education), and mother's country of birth (Sweden, other Scandinavian countries or others). As low birthweight might be an intermediate factor of the association between SDP and ADHD, we

choose not to adjust for that measure. Prepregnancy BMI is considered partly adjusted for by maternal education level and could potentially be an intermediate factor, thus not included as a covariate in the model.

### Statistical analyses

We used Cox proportional survival analysis to estimate the magnitude of the associations between SDP and offspring ADHD at the population level. The models calculated Hazard ratios (HRs) for time to ADHD diagnosis. Robust standard errors adjusted the 95% confidence intervals (CIs) for the presence of familial clustering in the analyses at the population level. In addition to a crude model, we also adjusted for the above-mentioned measured covariates.

To explore the effect of unmeasured familial confounding, we applied stratified Cox regression models to cousin and sibling data with a separate stratum for each set of full cousins and siblings, respectively. In the cousin sample, 41,028 cousins nested in 13,372 extended families were discordantly exposed to SDP, whereas in the sibling sample, there were 26,826 discordantly exposed siblings nested in 12,507 nuclear families. Cousin comparisons adjust for all unmeasured factors that are constant within the extended family, whereas the sibling comparisons adjust for factors that are constant within the nuclear family. The stratified Cox regression models using cousins were adjusted for the same covariates as in the models at the population level. The regression models applied to siblings were not adjusted for maternal highest education and mother's country of birth, as these were regarded constant or unlikely to change within siblings.

We conducted sensitivity analyses to assess the potential modifying effect of birth order on the studied associations. We also analyzed a birth year restricted sibling cohort to explore potential bias from outcome misclassification. Furthermore, we explored the generalizability of our results by comparing the sibling sample to the entire cohort (D'Onofrio et al., 2013). All statistical analyses were conducted in SAS software version 9.3 (SAS Institute, Cary, NC, USA).

## Results

Table 1 shows the distribution of offspring and maternal covariates. Offspring exposed to maternal SDP were more likely to be of late parity ( $p < .0001$ ), have mothers with lower education ( $p < .01$ ), and their mothers were less likely to cohabit with their biological father at childbirth.

In Table 2, the crude association showed that offspring exposed to maternal SDP were at increased risk for ADHD ( $HR_{ModerateSDP}$  1.89;  $HR_{HighSDP}$  2.50). This dose-dependent association decreased marginally after adjustment for measured covariates and

remained statistically significant ( $HR_{ModerateSDP}$  1.62;  $HR_{HighSDP}$  2.04). The associations were further attenuated in the cousin comparisons; that is, after adjustment for all unmeasured factors that are constant within extended families ( $HR_{ModerateSDP}$  1.45;  $HR_{HighSDP}$  1.69). Sibling comparisons showed that the associations observed at the population level were completely attenuated and no longer statistically significant ( $HR_{ModerateSDP}$  0.88;  $HR_{HighSDP}$  0.84), indicating that unmeasured familial factors that are constant within nuclear families explain the associations.

### Sensitivity analyses

Sensitivity analyses were conducted to explore assumptions and to test the robustness of the findings.

Second-born offspring were more often exposed to maternal SDP than first-born offspring. Therefore, we used another sibling sample including only first- and second-born offspring ( $N = 345,402$ ) from each nuclear family to test the modifying effect of birth order on the studied association. Stratified analyses at the population level showed similar results for first-born and second-born siblings ( $HR_{ModerateSDP\_FirstBorn}$  1.62 95% CI, 1.49–1.76;  $HR_{HighSDP\_FirstBorn}$  1.95 95% CI, 1.76–2.16;  $HR_{ModerateSDP\_SecondBorn}$  1.58 95% CI, 1.42–1.75; and  $HR_{HighSDP\_SecondBorn}$  2.13 95% CI, 1.90–2.39).

Because the outpatient register (started in 2001) and PDR (started in 2005) are relatively new, we explored potential bias from outcome misclassification. For instance, older siblings with an ADHD diagnosis early in life may not have been recorded with a diagnosis in this register. To minimize this potential bias, we performed a separate sibling comparison within a restricted cohort of offspring born 1994–1998. The two siblings from each nuclear family who were closest in age were selected for sibling comparisons, resulting in 123,974 full siblings nested within 61,987 families (the average age difference between siblings was 2.34 years). The results from the restricted sibling sample ( $HR_{ModerateSDP}$ , 1.01; 95% CI, 0.73–1.41;  $HR_{HighSDP}$  1.08; 95% CI, 0.70–1.66) were similar to that observed in the full sibling sample.

Finally, we explored the generalizability of the results from siblings to the entire cohort. We analyzed siblings as unrelated individuals using robust standard errors and the results in the sibling sample ( $HR_{ModerateSDP}$ , 1.62; 95% CI, 1.51–1.73;  $HR_{HighSDP}$  2.06; 95% CI, 1.90–2.22) were very similar to that in the entire cohort ( $HR_{ModerateSDP}$ , 1.62; 95% CI, 1.56–1.69;  $HR_{HighSDP}$  2.04; 95% CI, 1.95–2.13).

## Discussion

In this population-based cohort study, we used two different quasi-experimental designs (i.e., cousin



**Table 1** Demographic characteristics of mothers and offspring

	Entire cohort <sup>a</sup> (N = 768,227)	Cousins <sup>b</sup> (N = 155,852)	Full siblings <sup>c</sup> (N = 365,442)
Covariates	[N (%)]	[N (%)]	[N (%)]
SDP			
No cigarette	638,400 (83.1)	129,362 (83.0)	317,836 (87.0)
1–9 cigarettes per day	83,830 (10.9)	16,991 (10.9)	31,321 (8.6)
≥10 cigarettes per day	45,997 (6.0)	9,499 (6.1)	16,282 (4.5)
Sex			
Male	392,865 (51.1)	79,688 (51.1)	187,460 (51.3)
Female	375,282 (48.9)	76,164 (48.9)	177,982 (48.7)
Birth year			
1992–1994	301,273 (39.2)	60,917 (39.1)	124,104 (34.0)
1995–1997	247,989 (32.3)	51,532 (33.1)	142,840 (39.1)
1998–2000	218,965 (28.5)	43,403 (27.8)	98,498 (27.0)
Mother's parity			
1st	314,257 (40.9)	60,719 (39.0)	133,153 (36.4)
2nd	284,924 (37.1)	59,325 (38.1)	157,804 (43.2)
3rd	116,842 (15.2)	25,300 (16.2)	50,570 (13.8)
≥4th	52,187 (6.8)	10,508 (6.7)	23,915 (6.5)
Maternal age at childbirth (years)			
≤19	11,282 (1.5)	1,752 (1.1)	3,621 (0.99)
20–24	120,223 (15.6)	23,964 (15.4)	60,434 (16.5)
25–29	278,200 (36.2)	60,992 (39.1)	144,326 (39.5)
30–34	238,729 (31.1)	50,053 (32.1)	113,786 (31.0)
≥35	119,793 (15.6)	19,091 (12.3)	43,275 (11.8)
Cohabit with offspring's father			
Yes	697,508 (94.9)	148,971 (95.6)	355,678 (97.3)
No	37,309 (5.1)	6,881 (4.4)	9,768 (2.7)
Maternal highest education			
≤9 years	73,484 (9.8)	12,440 (8.0)	31,136 (8.7)
10–12 years	381,133 (50.6)	81,879 (52.5)	178,615 (49.8)
Graduate education	298,949 (39.7)	61,533 (39.5)	149,295 (41.6)
Mother's country of birth			
Sweden	647,746 (84.3)	148,511 (95.3)	310,254 (84.9)
Denmark, Finland, Norway, or Iceland	20,626 (2.7)	2,603 (1.7)	8,332 (2.3)
Other	99,837 (13.0)	4,738 (3.0)	46,851 (12.8)

<sup>a</sup>In the entire cohort 80 individuals were missing data for sex, 17 for mother's parity, 1 for maternal age at delivery, 33,410 for cohabitation status, 14,661 for maternal highest education, and 18 for mother's country of birth.

<sup>b</sup>Cousins with missing values on any of the above-mentioned covariates were precluded.

<sup>c</sup>In the full sibling sample, 6,396 individuals were missing data on education and 5 for mother's country of birth.

**Table 2** Relative risks of ADHD among offspring exposed to smoking during pregnancy

Exposure	HR (95% CI)			
	Crude <sup>a</sup>	Adjusted <sup>b</sup>	Cousins <sup>c</sup>	Full sibling <sup>d</sup>
No SDP	Reference	Reference	Reference	Reference
Moderate SDP (1–9 cigarettes per day)	1.89 (1.83–1.97)	1.62 (1.56–1.69)	1.45 (1.24–1.68)	0.88 (0.73–1.06)
High SDP (≥10 cigarettes per day)	2.50 (2.40–2.61)	2.04 (1.95–2.13)	1.69 (1.40–2.04)	0.84 (0.65–1.06)

<sup>a</sup>N = 768,227.

<sup>b</sup>N = 720,853 Adjusted for offspring gender, birth year, mother's parity, maternal age, cohabitation status, maternal highest education, and mother's country of birth.

<sup>c</sup>N = 155,852 Adjusted for offspring gender, birth year, mother's parity, maternal age, and cohabitation status.

<sup>d</sup>N = 365,442 Adjusted for offspring gender, birth year, mother's parity, maternal age, and cohabitation status.

and sibling comparisons) to explore the mechanisms through which maternal SDP influence ADHD. Our results suggest that the association between maternal SDP and offspring ADHD can be ascribed to unmeasured familial confounding. Similar results have been reported in quasi-experimental studies of ADHD from Finland (Obel et al., 2011), United Kingdom (Langley et al., 2012; Thapar et al., 2009),

and the United States (D'Onofrio et al., 2008), indicating that the study results generalize across samples with different demographic, racial, or ethnic characteristics. These results are also consistent with quasi-experimental studies of SDP for ADHD-related outcomes, such as disruptive behaviors, criminality, and academic/cognitive problems (D'Onofrio et al., 2013).

Our study addresses four critical limitations of previous sibling comparison studies (D'Onofrio et al., 2008; Lindblad & Hjern, 2010; Obel et al., 2011). First, our study was based on more than twice as many ADHD cases compared with the previous register-based sibling studies (Lindblad & Hjern, 2010; Obel et al., 2011), which allowed for more precise estimates (i.e., narrow confidence intervals) of the magnitude of the associations between SDP and offspring ADHD, even in the sibling and cousin comparisons. Second, the prior Swedish register-based study identified ADHD cases from medication use during one calendar year. This may have introduced bias in the sibling analyses due to misclassification of discordant and concordant sibling pairs. The large difference in the number of identified ADHD cases (despite similar N for the study population) between studies, coupled with sensitivity analyses based on a restricted cohort of siblings close in age, indicates that bias from outcome misclassification is unlikely in our study. Third, previous studies might be confounded by the fact that mothers reported on both SDP and offspring behaviors (D'Onofrio et al., 2008), while we in contrast used measured SDP via self-ratings at the first registration to antenatal care and obtained outcome information via clinical diagnosis made by physicians.

Finally, the results from previous sibling studies have been questioned by researchers who have explicitly hypothesized that women who vary in their smoking status across pregnancies are not comparable to all smoking women (Frisell, Oberg, Kuja-Halkola, & Sjolander, 2012). We used cousin comparisons to overcome the concern of external validity and observed converging evidence across different quasi-experimental designs, which further strengthen the familial confounding hypothesis.

Our results should be interpreted in the context of some limitations. Although sibling comparison will not be confounded by factors shared by siblings, the estimates might theoretically be more sensitive to bias due to nonshared confounders than the unpaired estimates (Frisell et al., 2012). The ascertainment of ADHD cases was predominantly based on ICD-10 diagnosis of hyperkinetic disorder and prescribed medication unique for the treatment of ADHD. The ICD-10 definition of ADHD is stricter compared with that in DSM-IV, and National guidelines for medication of ADHD, issued by the Swedish National Board of health and Welfare in 2002, stated that medication should be reserved for cases where other supportive interventions have failed, indicating that our proxies for ADHD most likely underestimate the incidence of ADHD and identify severe ADHD cases. Thus, our strategies probably could not avoid

producing false negatives, while we consider bias due to false positives more unlikely. It should also be noted that our findings do not rule out the possibility of an association between SDP and more refined ADHD-related neurocognitive deficits, such as executive function, response inhibition, and fine motor skills. To further explore this issue, studies using more detailed and refined outcome measures are warranted. As in all observational studies, we could not fully rule out residual confounding due to a lack of intact information on the exposure variable and other potential confounders. Also, as SDP was assessed at the first visit to antenatal care and used as a proxy for the entire pregnancy, there is a possibility of misclassification of exposure. In addition, pregnant women may also conceal their smoking habits (Lindqvist et al., 2002). However, previous studies have repeatedly shown support for a causal association between maternal SDP and low birthweight suggesting that the effect of exposure misclassification is small in magnitude (Cnattingius, 2004; D'Onofrio et al., 2012).

## Conclusion

In conclusion, our data suggest that the previously observed association between maternal SDP and ADHD can be ascribed to unmeasured familial confounding. There is mounting evidence that SDP is harmful in many ways (e.g., low birthweight and infant mortality) and our results should not be interpreted as an argument for changing the recommendations that women should not smoke during pregnancy. Nevertheless, we view our study as a contribution to the current state of knowledge concerning causal risk factors for ADHD. It is essential for clinicians, researchers, and policy-makers to focus on true causal risk factors and SDP is most probably not one of these.

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## Key points

- Maternal smoking during pregnancy is commonly cited in the literature as a risk factor for ADHD in offspring.
- In this nationwide register-based cohort study cousin and sibling data were used to control for unmeasured familial confounding.
- Our results suggest that the association between maternal smoking during pregnancy and offspring ADHD can be ascribed to unmeasured familial confounding.
- It is essential for clinicians, researchers, and policy-makers to focus on true causal risk factors, and maternal smoking during pregnancy is most probably not one of these.

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