

Growth of structures using redshift space distortion in $f(T)$ cosmology

Rebecca Briffa^{1,2}*, Celia Escamilla-Rivera,³ Jackson Levi Said^{1,2} and Jurgen Mifsud^{1,2}

¹*Institute of Space Sciences and Astronomy, University of Malta, Msida, MSD 2080, Malta*

²*Department of Physics, University of Malta, Msida, MSD 2080, Malta*

³*Instituto de Ciencias Nucleares, Universidad Nacional Autónoma de México, Circuito Exterior C.U., A.P. 70-543, México D.F. 04510, México*

Accepted 2024 January 8. Received 2023 December 19; in original form 2023 November 21

ABSTRACT

Cosmology faces a pressing challenge with the *Hubble* constant (H_0) tension, where the locally measured rate of the Universe's expansion does not align with predictions from the cosmic microwave background calibrated with lambda-cold dark matter model. Simultaneously, there is a growing tension involving the weighted amplitude of matter fluctuations, known as $S_{8,0}$ tension. Resolving both tensions within one framework would boost confidence in any one particular model. In this work, we analyse constraints in $f(T)$ gravity, a framework that shows promise in shedding light on cosmic evolution. We thoroughly examine prominent $f(T)$ gravity models using a combination of data sources, including pantheon+ (SN), cosmic chronometers, baryonic acoustic oscillations, and redshift space distortion data. We use these models to derive a spectrum of H_0 and $S_{8,0}$ values, aiming to gauge their ability to provide insights into, and potentially address, the challenges posed by the H_0 and $S_{8,0}$ tensions.

Key words: cosmological parameters – dark energy – large-scale structure of Universe.

1 INTRODUCTION

Measurements of the accelerating expansion of the Universe (Riess et al. 1998; Perlmutter et al. 1999) have led to the prospect that it may be expanding faster than predicted by the lambda-cold dark matter (Λ CDM) model (Di Valentino et al. 2021a). This may open the possibility that the concordance model description of gravitation through general relativity (GR), the as yet unobserved CDM (Gaitskell 2004; Bertone, Hooper & Silk 2005; Baudis 2016), and the theoretically problematic cosmological constant (Peebles & Ratra 2003; Copeland, Sami & Tsujikawa 2006) may require additions or corrections to its explanation of some kind. Over the last few years this has been expressed primarily through the value of the *Hubble* constant (Di Valentino et al. 2021c) and $f\sigma_{8,0}$ (Di Valentino et al. 2021b). The appearance of cosmic tensions has shaped into a tension between observations based on direct measurements of the late Universe such as those based on Type Ia supernovae (SNIa), the tip of the red giant branch measurements, strong lensing measurements (Freedman et al. 2020; Wong et al. 2020; Riess et al. 2022; Anderson, Koblishke & Eyer 2023), and indirect measurements coming from analyses of early time data from the cosmic microwave background (CMB) radiation as well as big bang nucleosynthesis (BBN) data (Cooke, Pettini & Steidel 2018; Alam et al. 2021; Abbott et al. 2022; Zhang et al. 2022) and others (Benisty, Davis & Evans 2023). Furthermore, recent data forecasting from new missions that include structure formation have increased this tension at local level (Atek et al. 2022; Lu et al. 2022), however, considering new systematic techniques on the measurements these tensions can be treated (Maldonado Alonso, Escamilla-Rivera & Sandoval-Orozco

2023). The growth in both these tensions has led to a reinvigorated exploration of possible modifications of the concordance model that have been developed in the literature over the last few decades.

The literature features a myriad of novel approaches in which to confront the growing issue of cosmic tensions. There have been proposals involving a re-evaluation of the cosmological principle (Krishnan et al. 2021, 2022), possible impacts of early Universe dark energy (Poulin, Smith & Karwal 2023), the appearance of extra degrees of freedom in the form of additional neutrino species in the early Universe (Di Valentino & Melchiorri 2022; Di Valentino et al. 2022), as well as modifications to gravity (Barack et al. 2019; Alves Batista et al. 2021; Saridakis et al. 2021; Abdalla et al. 2022) and others (Addazi et al. 2022). Many of these approaches modify a small part of the evolution of the Universe using new physics. On the other hand, modified gravity has the potential to make smaller changes that infiltrate the larger range of redshifts. Moreover, modifications to GR will also provide changes both at background and perturbative levels. In the space of possible modifications to GR, one approach that has become popular in recent years and which natively builds a competitive way in which to consider new physics is metric-affine gravity which is based on the exchange of the underlying connection with other possible geometries (Hehl et al. 1995; Beltrán Jiménez, Heisenberg & Koivisto 2019). In teleparallel gravity (TG), the geometric curvature of the Levi-Civita connection $\overset{\circ}{\Gamma}{}^{\sigma}{}_{\mu\nu}$ (overcircles denote quantities based on the curvature of the Levi-Civita connection in this work) is interchanged with the torsion of the teleparallel connection $\Gamma^{\sigma}{}_{\mu\nu}$. This may provide a more intuitive approach in which to consider new physics in gravitational theory.

The connection of TG is curvature-free (Cai et al. 2016; Krssak et al. 2019; Bahamonde et al. 2021) and produces an altogether novel formulation of gravity. There exists a particular combination of scalars in the theory that can produce a teleparallel equivalent of

* E-mail: rebecca.briffa.16@um.edu.mt

general relativity (TEGR), which is dynamically equivalent to GR but may have different IR completions (Mylova, Said & Saridakis 2023). A natural consequence of this is that both GR and TEGR are identical at the level of phenomenological predictions. Taking the same rationale as in $f(R)$ gravity (De Felice & Tsujikawa 2010; Sotiriou & Faraoni 2010; Capozziello & De Laurentis 2011), TEGR can be generalized to an $f(T)$ general class of models (Ferraro & Fiorini 2007, 2008; Bengochea & Ferraro 2009; Linder 2010; Chen et al. 2011; Farrugia & Levi Said 2016; Finch & Said 2018; Paliathanasis, Levi Said & Barrow 2018; Bahamonde, Flathmann & Pfeifer 2019; Bahamonde, Levi Said & Zubair 2020; Farrugia, Levi Said & Finch 2020; Bahamonde et al. 2022) where the TEGR Lagrangian is simple the torsion scalar T . This is an interesting model since it produces organically second-order equations of motion and agrees with the polarization modes of GR despite being fundamentally different.

$f(T)$ cosmology has been probed in various scenarios. At background level, $f(T)$ models have probed against several different types of data showing consistent results with the relatively high value of the *Hubble* constant (Briffa et al. 2022). Most recently this has been in work that incorporates the latest pantheon+ data set (Briffa et al. 2020; Cai, Khurshudyan & Saridakis 2020; Ren et al. 2022; Briffa et al. 2023). There has also been working that probes the early Universe using either BBN (Benetti, Capozziello & Lambiase 2020), or constraints from the CMB (Nunes 2018; Nunes, Pan & Saridakis 2018; Kumar, Nunes & Yadav 2023) with interesting results related to the best-fit values of the model parameters compared with both early and late time constraints. Also, constraints from primordial black holes seem to be consistent with a wide range of observations (Papanikolaou et al. 2023) giving more freedom to the possible cosmological models. From foundational physics recent work has also been done on non-flat cosmologies (Bahamonde et al. 2023) which may open a way to compare recent proposals in the literature (Di Valentino, Melchiorri & Silk 2019) on the topic in Λ CDM cosmology. Other works in the literature have explored possible effects on the fine structure constant (Levi Said et al. 2020) which are consistent with there being no variation.

In the current work, we are interested in determining constraints on models of $f(T)$ cosmology using $f\sigma_8(z)$ data. This window into the large-scale structure of the Universe will be a key tool in understanding the viability of new proposals for cosmology. The topic has been explored for previous data sets in Nesseris et al. (2013) where it was shown that $f(T)$ cosmological models are largely consistent with this probe of large-scale structure. Later in Anagnostopoulos, Basilakos & Saridakis (2019), these analyses were combined with background data sets which gave stricter constraints on model parameters. Recently, the $f\sigma_8(z)$ data have also been used to check for consistency with background data in the context of model-independent approaches to reconstructing modified cosmological models (Levi Said et al. 2021). We start by briefly introducing the background of TG and its predictions for $f\sigma_8$ in Section 2. We then discuss our observational data sets under consideration in Section 3. This is then followed by a presentation of our model constraints in Section 4, and a comparative analysis in Section 5. The results are summarized and discussed in our conclusion in Section 6.

2 $f(T)$ GRAVITY AND SCALAR PERTURBATIONS

TG is based on the exchange of curvature with torsion through the geometric connection, which is realized through the teleparallel connection $\Gamma_{\mu\nu}^\sigma$. This renders all measures of curvature identically

zero, since $\overset{\circ}{\Gamma}_{\mu\nu}^\sigma$ is curvature-less and satisfies metricity. While the regular Levi–Civita objects remain non-vanishing, such as the regular Riemann tensor $R_{\mu\nu\rho}^\sigma \neq 0$, its torsional analogue will vanish $R_{\mu\nu\rho}^\sigma = 0$. This means that a different class of torsional objects must be utilized to express the geometric deformation that embodies gravitation effects from the energy-momentum tensor (see reviews in Aldrovandi & Pereira 2013; Cai et al. 2016; Krssak et al. 2019; Bahamonde et al. 2021).

Another important aspect of TG is that it can be expressed in terms of a tetrad field that embodies the gravitational effect of the metric tensor through

$$g_{\mu\nu} = e_\mu^a e_\nu^b \eta_{ab}, \quad \eta_{ab} = E_A^\mu E_B^\nu g_{\mu\nu}, \quad (1)$$

where E_A^μ is the tetrad inverse, and where Latin indices represent local inertial coordinates, while Greek ones denote general space–time coordinates. For internal consistency, these tetrads must also satisfy orthogonality conditions

$$e_\mu^a E_B^\mu = \delta_b^a, \quad e_\mu^a E_A^\nu = \delta_\mu^a. \quad (2)$$

The tetrad encompasses the gravitational freedom of the metric, while the Lorentz invariance freedom can be assigned to the spin connection $\omega_{b\mu}^a$. The teleparallel connection can then be defined as Cai et al. (2016) and Krssak et al. (2019)

$$\Gamma_{\nu\mu}^\sigma := e_a^\sigma \partial_\mu e_\nu^a + E_A^\sigma \omega_{b\mu}^a e_\nu^b. \quad (3)$$

Given a metric, the infinite possible tetrad solutions in equation (1) are balanced by the spin connection components to retain general diffeomorphism invariance of the underlying theory. Together, the tetrad–spin connection pair defines a space–time.

The tensor structure of TG can be written in terms of the torsion tensor which is defined as

$$T_{\mu\nu}^\sigma := -2\Gamma_{[\mu\nu]}^\sigma, \quad (4)$$

where square brackets denote an antisymmetric operator which represents the translation field strength of TG (Aldrovandi & Pereira 2013). Taking a suitable choice of contractions of the torsion tensor, we can write the torsion scalar as Hayashi & Shirafuji (1979) and Bahamonde, Böhmer & Krššák (2017)

$$T := \frac{1}{4} T_{\mu\nu}^\alpha T_\alpha^{\mu\nu} + \frac{1}{2} T_{\mu\nu}^\alpha T_\alpha^{\nu\mu} - T_{\mu\alpha}^\nu T_\beta^{\beta\mu}, \quad (5)$$

which is dependent only on the teleparallel connection in an analogous way to the dependence of the Riemann tensor on the Levi–Civita connection.

Naturally, the teleparallel Ricci scalar will vanish, namely $R = 0$, while the teleparallel-based Ricci scalar turns out to be equal to the torsion scalar (up to a total divergence term), that is

$$R = \overset{\circ}{R} + T - B = 0, \quad (6)$$

where $B = (2/e)\partial_\mu(eT^\mu)$ is a boundary term and $e = \det(e_\mu^a) = \sqrt{-g}$ is the tetrad determinant. This relation ensures that the equations of motion of the Einstein–Hilbert action will be the same as those from the so-called TEGR which is based on the linear form of the torsion scalar. TEGR is dynamically equivalent to GR, but may differ in its UV completion. The TEGR Lagrangian can readily be generalized using the same reasoning as $f(R)$ gravity (De Felice & Tsujikawa 2010; Capozziello & De Laurentis 2011), except that we now consider an arbitrary function of the torsion scalar through (Ferraro & Fiorini 2007, 2008; Bengochea & Ferraro 2009; Linder 2010; Chen et al. 2011; Capozziello, De Falco & Ferrara 2022; Calzá

& Sebastiani 2023)

$$\begin{aligned} \mathcal{S}_{\mathcal{F}(T)} &= \frac{1}{2\kappa^2} \int d^4x e f(T) + \int d^4x e \mathcal{L}_m \\ &= \frac{1}{2\kappa^2} \int d^4x e (-T + \mathcal{F}(T)) + \int d^4x e \mathcal{L}_m, \end{aligned} \quad (7)$$

where $\kappa^2 = 8\pi G$, \mathcal{L}_m is the matter Lagrangian, and $e = \det(e_\mu^a) = \sqrt{-g}$ is the tetrad determinant.

The most impactful practical difference between $f(\overset{\circ}{R})$ and $\mathcal{F}(T)$ gravity theories is that the boundary term is no longer a total divergence term for $f(\overset{\circ}{R})$ gravity and so leads to fourth order equations of motion, while $\mathcal{F}(T)$ gravity retains generally second-order equations of motion. This may be advantageous both for numerical reasons as well as theoretical motivations when considering certain types of ghosts in the theory. By taking a variation with respect to the tetrad field, the field equations follow

$$\begin{aligned} W_\rho^\mu &:= e^{-1} \partial_\nu (e E_A^\rho S^{\mu\nu}) (-1 + \mathcal{F}_T) - E_A^\lambda T_{\nu\lambda}^\rho S^{\nu\mu} (-1 + \mathcal{F}_T) \\ &+ \frac{1}{4} E_A^\mu (-T + \mathcal{F}(T)) + E_A^\rho S^{\mu\nu} \partial_\nu (T) \mathcal{F}_{TT} \\ &+ E_B^\lambda \omega_{ab}^b S_\lambda^{\nu\mu} (-1 + \mathcal{F}_T) = \kappa^2 E_A^\rho \Theta_\rho^\mu, \end{aligned} \quad (8)$$

where subscripts denote derivatives ($\mathcal{F}_T = \partial\mathcal{F}/\partial T$ and $\mathcal{F}_{TT} = \partial^2\mathcal{F}/\partial T^2$), and $\Theta_\rho^\nu = \delta\mathcal{L}_m/\delta e_\mu^\rho$ is the regular energy-momentum tensor. This limits to TEGR as $\mathcal{F}(T) \rightarrow 0$. TG contains two dynamical variables, namely, the tetrad and spin connection which both have associated variations of the $\mathcal{F}(T)$ action. It turns out that they can both be expressed in terms of the above tetrad variation as

$$W_{(\mu\nu)} = \kappa^2 \Theta_{\mu\nu}, \quad \text{and} \quad W_{[\mu\nu]} = 0, \quad (9)$$

where the tetrad and spin connection field equations are, respectively, the symmetric and antisymmetric operators acted upon $W_{\mu\nu}$ (Bahamonde et al. 2021). Indeed, a particular choice of tetrad components exists in which the spin connection components vanish, called the Weitzenböck gauge. Here, the spin connection equations on motion are identically satisfied. More generally, the six spin connection field equations express the Lorentz invariance freedom (three translations and three rotations), while the 10 tetrad field equations manifest the gravitational equations of motion.

We can now consider a flat homogeneous and isotropic cosmology through the regular flat Friedmann–Lemaître–Robertson–Walker (FLRW) metric

$$ds^2 = -dt^2 + a^2(t) (dx^2 + dy^2 + dz^2), \quad (10)$$

from which we can identify the *Hubble* parameter as $H = \dot{a}/a$, and where overdots refer to derivatives with respect to cosmic time. It can be shown that the tetrad (Tamanini & Boehmer 2012; Krššák & Saridakis 2016)

$$e_\mu^a = \text{diag}(1, a(t), a(t), a(t)), \quad (11)$$

is compatible with the Weitzenböck gauge and so the spin connection components can be set to zero (Hohmann et al. 2019). By considering the tetrad field equations (9), the modified Friedmann equations can be written as

$$H^2 - \frac{T}{3} \mathcal{F}_T + \frac{\mathcal{F}}{6} = \frac{\kappa^2}{3} \rho, \quad (12)$$

$$\dot{H} (1 - \mathcal{F}_T - 2T \mathcal{F}_{TT}) = -\frac{\kappa^2}{2} (\rho + p), \quad (13)$$

where ρ and p denote the energy density and pressure of the matter components, respectively.

TG can be used to interpret large-scale structure data by considering scalar perturbations of the flat FLRW metric. We probe this data

through the growth rate measurements of $f\sigma_8(z)$ from RSD. This is expressed by the logarithm derivative of the matter perturbation $\delta(z)$ with respect to the logarithm of the cosmic scale factor, namely

$$f(z) = \frac{d \ln \delta(z)}{d \ln a} = -(1+z) \frac{d \ln \delta(z)}{dz} = -(1+z) \frac{\delta'(z)}{\delta(z)}, \quad (14)$$

where a prime denotes a derivative with respect to redshift $z = a^{-1} - 1$. On the other hand, the linear theory root-mean-square mass fluctuation within a sphere of radius $8 h^{-1}$ Mpc can be expressed as

$$\sigma_8(z) = \sigma_{8,0} \frac{\delta(z)}{\delta_0}, \quad (15)$$

where a 0-subscript denotes the respective value at $z = 0$. Thus, the growth rate can be generally written as

$$f\sigma_8(z) = -(1+z) \frac{\sigma_{8,0}}{\delta_0} \delta'(z), \quad (16)$$

which directly leads to the normalized form of $\delta'(z)$ through

$$\frac{\delta'(z)}{\delta_0} = -\frac{1}{\sigma_{8,0}} \frac{f\sigma_8(z)}{(1+z)}. \quad (17)$$

By integrating this expression, the normalized redshift evolution of the matter perturbation can be written as

$$\frac{\delta(z)}{\delta_0} = 1 - \frac{1}{\sigma_{8,0}} \int_0^z \frac{f\sigma_8(\tilde{z})}{(1+\tilde{z})} d\tilde{z}, \quad (18)$$

while equation (17) also gives the second-order derivative

$$\frac{\delta''(z)}{\delta_0} = -\frac{1}{\sigma_{8,0}} \left[\frac{(1+z) f\sigma_8'(z) - f\sigma_8(z)}{(1+z)^2} \right]. \quad (19)$$

The result is the redshift evolution of f can now be determined

$$\begin{aligned} f'(z) &= -\left(\frac{\delta'(z)}{\delta_0} \right) \left(\frac{\delta_0}{\delta(z)} \right) \\ &- (1+z) \left[\left(\frac{\delta''(z)}{\delta_0} \right) \left(\frac{\delta_0}{\delta(z)} \right) - \left(\frac{\delta'(z)}{\delta_0} \right)^2 \left(\frac{\delta_0}{\delta(z)} \right)^2 \right]. \end{aligned} \quad (20)$$

In the subhorizon limit, the linear matter perturbations equation can then be written as

$$\ddot{\delta} + 2H\dot{\delta} = 4\pi G_{\text{eff}} \rho \delta, \quad (21)$$

where G_{eff} is the effective Newton's constant which is in general a function of z and cosmic wave vector \mathbf{k} (Amendola & Tsujikawa 2010; Hernández 2017). However for this limit and the data sets under consideration, the G_{eff} can be taken to be independent of \mathbf{k} . In this setting, equation (22) takes the form

$$\delta''(z) + \left(\frac{H'(z)}{H(z)} - \frac{1}{1+z} \right) \delta'(z) = \frac{3}{2} \frac{G_{\text{eff}}(z)}{G_N} \left(\frac{H_0}{H(z)} \right)^2 \Omega_{m,0} (1+z) \delta(z), \quad (22)$$

where $\Omega_{m,0}$ denotes the current matter fractional density, H_0 is *Hubble's* constant, and G_N is Newton's gravitational constant. We can also express this relation in terms of the growth rate f , giving

$$\begin{aligned} f^2(z) + \left[2 - (1+z) \frac{H'(z)}{H(z)} \right] f(z) - (1+z) f'(z) \\ = \frac{3}{2} \frac{G_{\text{eff}}(z)}{G_N} \left(\frac{H_0}{H(z)} \right)^2 \Omega_{m,0} (1+z)^3. \end{aligned} \quad (23)$$

In the case of $\mathcal{F}(T)$ gravity, the linear matter perturbation evolution equation (18) is expressed through Golovnev & Koivisto (2018), Nunes (2018), Levi Said et al. (2020), and Sahu et al. (2020)

$$G_{\text{eff}}(z) = \frac{G_N}{1 + \mathcal{F}_T(z)}. \quad (24)$$

Hence, the linear matter perturbation equation is sensitive to $\mathcal{F}(T)$ gravity and so large-scale structure data sets can be used to probe observational constraints on potential models.

In order to utilize equations like equations (22) and (23), we need to determine the value of $H'(z)$, obtained from the second Friedmann equation (13), via

$$H'(z) = \frac{\kappa^2(p + \rho)}{2H(1+z)(1 - \mathcal{F}_T - 2T\mathcal{F}_{TT})}. \quad (25)$$

3 OBSERVATIONAL DATA

In this section, we will provide an overview of how we utilized observational data in our research to evaluate the most favourable $f(T)$ models from the literature. We include a variety of observational data, such as the pantheon+ sample of SNIa data, baryon acoustic oscillations (BAO) data, cosmic chronometers (CC) data, and growth rate (redshift space distortion, RSD) data. To analyse the data, we implemented an MCMC (Monte Carlo Markov Chain) approach using the EMCEE package publicly available at Foreman-Mackey et al. (2013). This enabled us to constrain the model and cosmological parameters and thus, investigate the posterior of the parameter space. This yielded one-dimensional distributions that illustrate the parameter's posterior distribution, and two-dimensional distributions that demonstrate the covariance between two different parameters. These distributions were complemented by the 1σ and 2σ confidence levels as will be shown in Section 4.

3.1 Cosmic chronometers

CC offer a useful tool to directly constrain the *Hubble* rate $H(z)$ at different redshifts. To this end, we adopt 31 data points as compiled from Jimenez et al. (2003), Simon, Verde & Jimenez (2005), Stern et al. (2010), Moresco et al. (2012, 2016), Zhang et al. (2014), and Moresco (2015). The CC method involves using spectroscopic dating techniques on passively evolving galaxies to estimate the age difference between two galaxies at different redshifts. By measuring this age difference, $\Delta z/\Delta t$ can be inferred from observations, allowing for the computation of $H(z) = -(1+z)^{-1}\Delta z/\Delta t$. Thus, CC data were found to be particularly reliable because they are independent of any cosmological model, avoid any complex integration, and rely on the absolute age determination of galaxies (Jimenez & Loeb 2002).

In our MCMC analysis, we used χ_{CC}^2 to evaluate the agreement between the theoretical *Hubble* parameter values $H(z_i, \Theta)$, with model parameters Θ , and the observational *Hubble* data values $H_{\text{obs}}(z_i)$, with an observational error of $\sigma_H(z_i)$. The χ_{CC}^2 was calculated using

$$\chi_{CC}^2 = \sum_{i=1}^{39} \frac{(H(z_i, \Theta) - H_{\text{obs}}(z_i))^2}{\sigma_H^2(z_i)}. \quad (26)$$

3.2 Type Ia supernovae–pantheon+ data set

The SNIa data set used in this study is the pantheon+ (PN+ & SHOES) sample (Scolnic et al. 2022), which is one of the largest compilations of SNIa data and contains 1701 SNIa measurements spanning a redshift range of 0.0–2.26. The pantheon+ analysis (Brout et al. 2022), incorporates SHOES Cepheid host distance anchors, Riess et al. (2019) in the likelihood, effectively addressing the degeneracy between parameters M and H_0 in SNIa analysis. Following this approach, we also adopt the label ‘PN+ & SHOES’ to signify our incorporation of the same methodology.

SNIa can be used to determine the expansion rate of the Universe, $H(z)$, by measuring the observed apparent magnitude, m , and the absolute magnitude, M , through the equation:

$$\mu(z_i, \Theta) = m - M = 5 \log_{10}[D_L(z_i, \Theta)] + 25, \quad (27)$$

where Θ represents the set of cosmological parameters that describe the Universe, z_i is the redshift of the SNIa measurement, and $D_L(z_i, \Theta)$ is the luminosity distance given by

$$D_L(z_i, \Theta) = c(1+z_i) \int_0^{z_i} \frac{dz'}{H(z', \Theta)}, \quad (28)$$

where c is the speed of light. To calibrate the observed apparent magnitude of each SNIa, a fiducial absolute magnitude M is used, as shown in equation (27). In our MCMC analyses, we treat M as a nuisance parameter.

To constrain the cosmological parameters, the minimum χ^2 is calculated through Conley et al. (2011),

$$\chi_{SN}^2 = (\Delta\mu(z_i, \Theta))^T C_{SN}^{-1} (\Delta\mu(z_i, \Theta)), \quad (29)$$

where $(\Delta\mu(z_i, \Theta)) = (\mu(z_i, \Theta) - \mu(z_i)_{\text{obs}})$ and C_{SN} is the corresponding covariance matrix which accounts for the statistical and systematic uncertainties.

3.3 Baryon acoustic oscillations

In this study, we use a variety of observational missions to constrain cosmological parameters. These data sets include measurements from the 6dF Galaxy Survey at an effective redshift of $z_{\text{eff}} = 0.106$ (Beutler et al. 2011), the Baryon Oscillation Spectroscopic Survey (BOSS) Data Release 11 (DR11) quasar Lyman-alpha measurements at an effective redshift of $z_{\text{eff}} = 2.4$ (du Mas des Bourboux et al. 2017), and the Sloan Digital Sky Survey (SDSS) Main Galaxy sample at an effective redshift of $z_{\text{eff}} = 0.15$ (Ross et al. 2015). Additionally, measurements of the *Hubble* parameter and the corresponding comoving angular diameter at $z_{\text{eff}} = 0.38, 0.51$ were obtained from the third generation of the SDSS mission (SDSS BOSS DR12, Alam et al. 2021). We also include $H(z)$ measurements and angular diameter distances at $z_{\text{eff}} = 0.98, 1.23, 1.52, \text{ and } 1.94$ from the fourth generation of the SDSS mission (SDSS-IV BOSS DR12, Zhao et al. 2019).

As has been mentioned these different data sets report different observational quantities that are related to one another. For the BAO data sets under consideration, we compute the *Hubble* distance $D_H(z)$ given by $D_H(z) = \frac{c}{H(z)}$. We also consider the angular diameter distance $D_A(z)$ defined as

$$D_A(z) = \frac{c}{1+z} \int_0^z \frac{dz'}{H(z')}, \quad (30)$$

from which two other quantities can be derived. The first is the comoving angular diameter distance D_M given trough $D_M = (1+z)D_A(z)$ whilst the second one is the volume average distance given by

$$D_V(z) = (1+z)^2 \left[D_A(z)^2 \frac{cz}{H(z)} \right]^{\frac{1}{3}}. \quad (31)$$

Using the BAO results, we calculate the corresponding combination of results $\mathcal{G}(z_i) = D_V(z_i)/r_s(z_d)$, $r_s(z_d)/D_V(z_i)$, $D_H(z_i)$, $D_M(z_i)(r_{s,\text{fid}}(z_d)/r_s(z_d))$, $H(z_i)(r_s(z_d)/r_{s,\text{fid}}(z_d))$, $D_A(z_i)(r_{s,\text{fid}}(z_d)/r_s(z_d))$. In this case, we require the comoving sound horizon at the end of the baryon drag epoch at $z_d \sim 1059.94$ (Aghanim et al. 2020) which can be calculated using

$$\begin{aligned} r_s(z) &= \int_z^\infty \frac{c_s(\bar{z})}{H(\bar{z})} d\bar{z} \\ &= \frac{1}{\sqrt{3}} \int_0^{1/(1+z)} \frac{da}{a^2 H(a) \sqrt{1 + [3\Omega_{b,0}/(4\Omega_{\gamma,0})] a}}, \end{aligned} \quad (32)$$

where we have adopted a fiducial value of $r_{s, \text{fid}}(z_d) = 147.78 \text{ Mpc}$ (Aghanim et al. 2020) with an assumption of $\Omega_{b,0} = 0.02242$ (Aghanim et al. 2020) and $T_0 = 2.7255 \text{ K}$ (Fixsen 2009). Therefore, the corresponding χ^2 is calculated through

$$\chi_{\text{BAO}}^2(\Theta) = \Delta G(z_i, \Theta)^T C_{\text{BAO}}^{-1} \Delta G(z_i, \Theta), \quad (33)$$

where $\Delta G(z_i, \Theta) = G(z_i, \Theta) - G_{\text{obs}}(z_i)$ and C_{BAO} is the corresponding covariance matrix for the BAO observations. The total χ_{BAO}^2 is therefore the sum of all the individual χ^2 corresponding to each data set.

3.4 Growth rate data (RSD)

The final data compilation we used in this work is the growth rate data compilation presented in table II of appendix A in Alestas, Kazantzidis & Nesseris (2022). This data set consists of measurements of the growth of cosmic structure and is commonly referred to as RSD data, due to a phenomenon that occurs during observations at both large and small scales. The peculiar velocity of galaxies causes high-density regions of the Universe to appear elongated in the line-of-sight direction at small scales and flattened at large scales. As a result, maps of galaxies where distances are measured from spectroscopic redshifts exhibit anisotropic deviations from the true galaxy distribution. Therefore, RSD measurements can provide valuable insights into the large-scale structure of the Universe, which is shaped by the underlying theory of gravity governing the evolution and formation of cosmic structure. Consequently, RSD presents a promising approach to testing modified theories of gravity.

The growth rate $f(z)$, can be estimated using RSD cosmological probes as a way to constrain cosmological models (Gupta, Sen & Sen 2012; Gonzalez, Alcaniz & Carvalho 2016; Lambiase et al. 2019). However, instead of reporting the growth rate directly, Large Scale Structure (LSS) surveys typically report the density-weighted growth rate, $f\sigma_8 \equiv f(z)\sigma_8(z)$ which is bias-independent (Kazantzidis & Perivolaropoulos 2018). Furthermore, we should acknowledge the Alcock–Paczynski effect, which emerges from the requirement of assuming a cosmological model for the conversion of redshifts into distances. However, recent studies, including those outlined in Levi Said et al. (2021), have indicated that the influence of this effect is minimal and thus has been disregarded in our analysis.

Therefore, while the theoretical prediction can be made from equation (16), the aforementioned RSD data can be used to constrain cosmological parameters, in particular $\sigma_{8,0}$ such that the corresponding χ_{min}^2 can be given as

$$\chi_{\text{RSD}}^2 = \Delta Q(z_i, \Theta)^T C_{\text{RSD}}^{-1} \Delta Q(z_i, \Theta), \quad (34)$$

where $Q(z_i, \Theta) = (f\sigma_8(z_i, \Theta)_{\text{theo}} - f\sigma_{8, \text{obs}}(z_i))$ and C_{RSD}^{-1} is the inverse covariance matrix which is assumed to be a diagonal matrix except for the WiggleZ subset data which can be found in Blake et al. (2012). Therefore, the total covariance matrix can be written as

$$C_{\text{RSD}} = \begin{pmatrix} \sigma_1^2 & 0 & 0 & \dots \\ 0 & C_{\text{WiggleZ}} & 0 & \dots \\ 0 & 0 & \dots & \sigma_N^2 \end{pmatrix}. \quad (35)$$

To compute $Q(z_i, \Theta)$, we start by obtaining the theoretical values of $f\sigma_8(z_i, \Theta)_{\text{theo}}$ using equation (22) along with equation (16). On the other hand, $f\sigma_8(z_i, \Theta)_{\text{obs}}$ is extracted from the compiled observational data as described earlier.

4 RESULTS

In this section, we delve into the outcomes yielded by various combinations of data sets. Each subsection is dedicated to distinct $f(T)$ models, selected based on their potential to accurately reflect our cosmological history, as demonstrated in previous studies (Nunes et al. 2016; Briffa et al. 2022, 2023; dos Santos, Gonzalez & Silva 2022; Sandoval-Orozco et al. 2023). For every model, we showcase the 1σ and 2σ uncertainty ranges. These results are accompanied by a table showing the final results that include the Hubble constant (H_0) expressed in $\text{kms}^{-1}\text{Mpc}^{-1}$, the matter density parameter $\Omega_{m,0}$ and $\sigma_{8,0}$ along with other model parameters. This will give us an insight into how the different models and data set combinations affect the H_0 tension and $S_{8,0}$ cosmological tensions.

4.1 Power-law model

The power-law model, initially proposed by Bengochea and Ferraro in their work (Bengochea & Ferraro 2009), was adopted due to its capability to replicate the observed late-universe acceleration without the need for dark energy. Henceforth, this model will be referred to as $f_1\text{CDM}$ and is specified by

$$\mathcal{F}_1(T) = \alpha_1(-T)^{p_1}. \quad (36)$$

This power-law form is characterized by two constants, namely α_1 and p_1 which is introduced as a modification to $\mathcal{F}(T)$ function. Being a constant α_1 can be evaluated at any time and therefore, it can be calculated at the current time by using the Friedmann equation (12), which results in

$$\alpha_1 = (6H_0^2)^{1-p_1} \frac{1 - \Omega_{m,0} - \Omega_{r,0}}{1 - 2p_1}, \quad (37)$$

where the density parameter for matter and radiation is represented by $\Omega_{m,0}$ and $\Omega_{r,0}$ respectively. Consequently, by utilizing equation (37), the $f_1\text{CDM}$ model introduces only one new parameter, namely p_1 , thereby enhancing its elegance and simplicity. In combination with the parameters H_0 , $\Omega_{m,0}$, and $\sigma_{8,0}$, we determine the model parameter p_1 through MCMC analyses performed on observational data.

By substituting the above equation into the Friedmann equation (12), we derive the resulting Friedmann equation for this model

$$E^2(z) = \Omega_{m,0}(1+z)^3 + \Omega_{r,0}(1+z)^4 + (1 - \Omega_{m,0} - \Omega_{r,0})E^{2p_1}(z), \quad (38)$$

where $E(z) := \frac{H(z)}{H_0}$. It is worth noting that this equation cannot be solved analytically, so we employ numerical methods to calculate $E(z)$ at each redshift point. Consequently, using MCMC analysis, we solve for each redshift point where observational data exists to obtain parameter values. Specifically, we extract values for H_0 , $\Omega_{m,0}$, and p_1 . Additionally, when RSD data are included, we can obtain values for $\sigma_{8,0}$ allowing us to assess the model's compatibility with observational data and refine our understanding of the Universe's properties. When examining equation (36), it becomes apparent that the ΛCDM limit can be attained when $p_1 = 0$. On the other hand, when $p_1 = 1$, the model converges to the GR limit. In this scenario, an additional component in the Friedmann equation emerges, introducing a rescaled gravitational constant term within the density parameters.

The contour plots together with their posteriors for the $f_1\text{CDM}$ model are shown in Fig. 1. In this and subsequent models, the blue and green contours represent the CC+BAO and CC+PN⁺+BAO data sets, respectively. The remaining colours correspond to various data sets that incorporate the RSD data. On the other hand, the

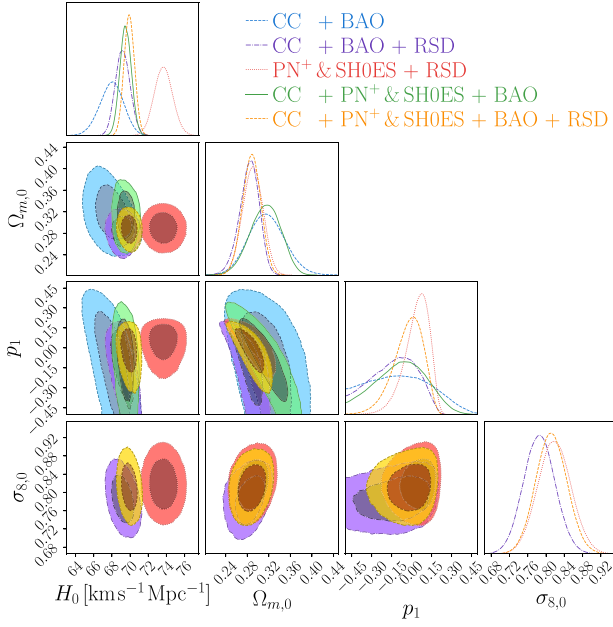


Figure 1. Confidence contours and posterior distributions for the f_1 CDM model (power-law model) parameters, including H_0 , $\Omega_{m,0}$, and p_1 . In cases where the RSD data are incorporated (purple, red, and yellow contours), the $\sigma_{8,0}$ parameter is also displayed.

precise values for the parameters are tabulated in Table 1. Notably, this table highlights that the highest H_0 value is obtained when combining the PN⁺ & SH0ES data with the RSD data is used. This result aligns with expectations, as the PN⁺ & SH0ES data converges with values reported by the SH0ES team, which reports a value of $H_0 = 73.30 \pm 1.04 \text{ km s}^{-1} \text{ Mpc}^{-1}$ (Scolnic et al. 2022). Conversely, the lowest value for the H_0 parameter is derived from the CC+BAO data sets, primarily influenced by the BAO data, which relates to conditions in the early Universe.

It is noteworthy to notice the relationship between the *Hubble* parameter (H_0) and the matter density parameter ($\Omega_{m,0}$). As H_0 increases, indicating a faster expansion rate of the Universe, the matter density decreases. This effect is particularly pronounced in the PN⁺ & SH0ES+RSD combination, where the Universe’s energy predominantly manifests as an effective form of dark energy, driven by the elevated H_0 parameter. The presence of this phenomenon is further confirmed by the apparent anticorrelation depicted in Fig. 1.

As noted, the Λ CDM limit is reached for when $p_1 = 0$. Indeed, the resultant values for p_1 are in proximity of 0, and this limit falls within the 1σ range. Furthermore, it can be observed that the inclusion of the RSD data leads to tighter constraints on the p_1 parameter. However, this effect is not exclusive to the p_1 parameter alone; it also extends to other parameters, as evident from the contour plots in Fig. 1.

Table 1. Exact results for f_1 model that include the parameters H_0 , $\Omega_{m,0}$, and p_1 . The $\sigma_{8,0}$ parameter and the nuisance parameter M , are provided for data sets that include RSD or PN⁺ & SH0ES, respectively otherwise, they are left empty.

Data sets	H_0 [km s ⁻¹ Mpc ⁻¹]	$\Omega_{m,0}$	p_1	$\sigma_{8,0}$	M
CC + BAO	$68.1^{+1.2}_{-1.4}$	$0.314^{+0.034}_{-0.033}$	$-0.09^{+0.24}_{-0.30}$	–	–
CC + BAO + RSD	69.17 ± 0.81	$0.287^{+0.016}_{-0.020}$	$-0.09^{+0.17}_{-0.20}$	0.785 ± 0.035	–
PN ⁺ & SH0ES + RSD	73.7 ± 1.0	$0.290^{+0.019}_{-0.018}$	$0.076^{+0.075}_{-0.102}$	$0.817^{+0.037}_{-0.035}$	$-19.252^{+0.029}_{-0.030}$
CC + PN ⁺ & SH0ES + BAO	$69.45^{+0.69}_{-0.58}$	$0.316^{+0.028}_{-0.029}$	$-0.06^{+0.19}_{-0.22}$	–	-19.375 ± 0.017
CC + PN ⁺ & SH0ES + BAO + RSD	69.90 ± 0.58	$0.289^{+0.016}_{-0.018}$	$0.014^{+0.091}_{-0.125}$	$0.810^{+0.036}_{-0.033}$	$-19.367^{+0.016}_{-0.017}$

Incorporating the RSD data has been instrumental in constraining the amplitude of mass fluctuations, $\sigma_{8,0}$. Upon initial examination of Table 1, it appears that there is a correlation between the H_0 parameter and $\sigma_{8,0}$, where a higher H_0 tends to correspond to a slightly higher $\sigma_{8,0}$. However, the contour plots suggest a more complex and degenerate relationship between these parameters. It is worth noting that the RSD data appears to have more Gaussian errors when compared to the other data sets. The influence of these growth structure data is encapsulated by the parameter G_{eff} , which, in this model, is expressed as

$$G_{\text{eff}} = \frac{G_N}{1 - \alpha_1 p_1 (-T)^{p_1 - 1}}, \quad (39)$$

where the specific values for the parameters are extracted from the relevant table, and as a result, we observe that G_{eff} approximates G_N under these conditions.

Additionally, the constraint on these parameters has enabled us to explore a tension quantified in terms of $S_{8,0} \equiv \sigma_{8,0} \sqrt{\Omega_{m,0}/0.3}$ (Anchordoqui et al. 2021; Beltrán Jiménez et al. 2021; Benisty 2021; Brieden, Gil-Marín & Verde 2023; Clark et al. 2023; Preston, Amon & Efstathiou 2023; Rubira, Mazoun & Garny 2023). The results and the posteriors of these parameters can be found in Table 2 and Fig. 2, respectively. Reflecting the values obtained for $\sigma_{8,0}$, the highest value observed for $S_{8,0}$ was attained for the PN⁺ & SH0ES+RSD combination, measuring at $S_{8,0} = 0.801^{+0.052}_{-0.046}$. However, it is worth noting that we also provide values for the RSD data set alone to isolate the impact of RSD data on this parameter, where in this case, the value for the RSD data set reaches a minimum. Furthermore, it is evident that across all data sets, the constraints for such parameters are notably tight. We also examine the relationship between the $S_{8,0}$ parameter and the model parameter p_1 , which is depicted in Fig. A1 found in Appendix A. A significant anticorrelation exists between these two parameters. In return, this might suggest that changes in the power-law exponent (p_1) might have a direct impact on the amplitude of mass fluctuations ($S_{8,0}$).

4.2 Linder model

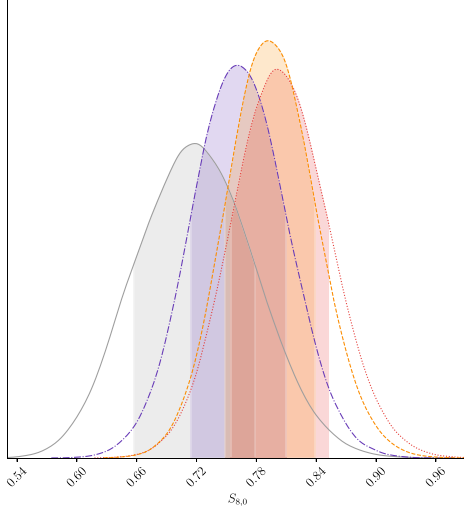
The second model under consideration is the Linder (2009) model, explicitly developed to explain the Universe’s late-time acceleration without invoking the presence of dark energy. This model introduces an exponential component that incorporates the torsion scalar, denoted as T , and is expressed as follows

$$\mathcal{F}_2 = \alpha_2 T_0 \left(1 - \text{Exp} \left[-p_2 \sqrt{T/T_0} \right] \right), \quad (40)$$

where both α_2 and p_2 are constants, while $T_0 = T|_{t=t_0} = -6H_0^2$ represents the torsion scalar at the current epoch. Similar to the previous model, if α_2 is a constant, it can be immediately determined from the Friedman equation at the present time,

Table 2. Exact $S_{8,0}$ values corresponding to various data sets for the f_1 CDM model.

Data sets	$S_{8,0}$
— RSD	$0.718^{+0.061}_{-0.060}$
--- CC + BAO + RSD	$0.761^{+0.049}_{-0.046}$
⋯⋯ PN ⁺ & SH0ES + RSD	$0.801^{+0.052}_{-0.046}$
- - - CC + PN ⁺ & SH0ES + BAO + RSD	$0.792^{+0.047}_{-0.043}$


Figure 2. Posterior distribution for the $S_{8,0}$ parameter in the f_1 CDM model. Legend: grey denotes the RSD data, purple corresponds to CC+BAO+RSD data, red represents the PN⁺ & SH0ES + RSD data set, while orange indicates CC + PN⁺ & SH0ES + BAO + RSD data.

yielding

$$\alpha_2 = \frac{1 - \Omega_{m,0} - \Omega_{r,0}}{(1 + p_2)e^{-p_2} - 1}. \quad (41)$$

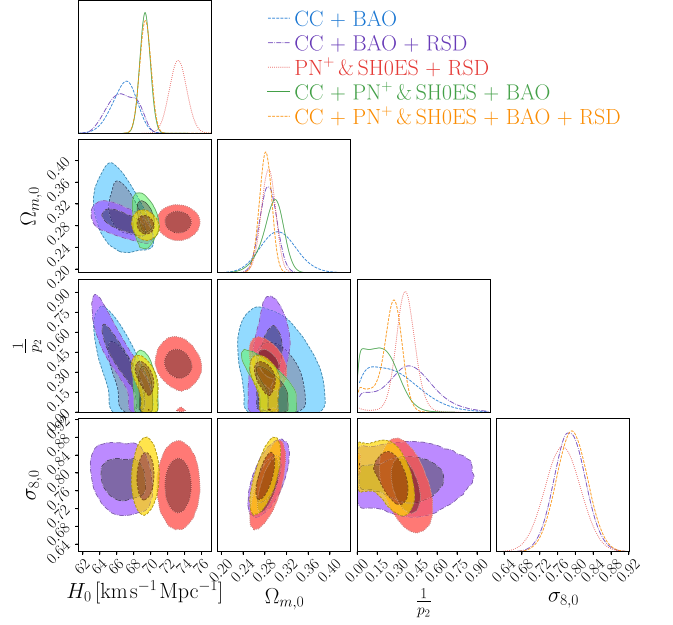
Therefore, in the Linder model, hereafter referred to as f_2 CDM, the only new model parameter is denoted as p_2 , and it will be constrained through the MCMC analysis. Consequently, equation (41) allows us to express the Friedmann equation in terms of H_0 , $\Omega_{m,0}$, and p_2 , resulting in the following form

$$E^2(z) = \Omega_{m,0}(1+z)^3 + \Omega_{r,0}(1+z)^4 + \frac{1 - \Omega_{m,0} - \Omega_{r,0}}{(p_2 + 1)e^{-p_2} - 1} \times [(1 + p_2 E(z)) \text{Exp}[-p_2 E(z)] - 1]. \quad (42)$$

Contrary to the f_1 CDM model, this model can be reduced to the Λ CDM when $p_2 \rightarrow \infty$. However, for numerical stability, the analysis

Table 3. Exact results for f_2 model that include the parameters H_0 , $\Omega_{m,0}$, and $\frac{1}{p_2}$. The $\sigma_{8,0}$ parameter and the nuisance parameter M , are provided for data sets that include RSD or PN⁺ & SH0ES, respectively otherwise, they are left empty.

Data Sets	H_0 [km s ⁻¹ Mpc ⁻¹]	$\Omega_{m,0}$	$\frac{1}{p_2}$	$\sigma_{8,0}$	M
CC + BAO	$67.2^{+1.2}_{-1.6}$	$0.302^{+0.035}_{-0.030}$	$0.00^{+0.37}_{-0.00}$	—	—
CC + BAO + RSD	$66.5^{+2.2}_{-1.3}$	$0.286^{+0.016}_{-0.015}$	$0.39^{+0.21}_{-0.22}$	$0.784^{+0.038}_{-0.032}$	—
PN ⁺ & SH0ES + RSD	73.2 ± 1.0	0.287 ± 0.013	$0.359^{+0.077}_{-0.071}$	$0.770^{+0.042}_{-0.039}$	$-19.26^{+0.33}_{-0.31}$
CC + PN ⁺ & SH0ES + BAO	$69.35^{+0.61}_{-0.63}$	$0.299^{+0.017}_{-0.021}$	$0.167^{+0.080}_{-0.154}$	—	$-19.40^{+0.21}_{-0.16}$
CC + PN ⁺ & SH0ES + BAO + RSD	$69.38^{+0.67}_{-0.68}$	0.282 ± 0.011	$0.275^{+0.083}_{-0.096}$	0.793 ± 0.035	$-19.37^{+0.31}_{-0.30}$


Figure 3. Confidence contours and posterior distributions for the f_2 CDM model (Linder model) parameters, including H_0 , $\Omega_{m,0}$, and $\frac{1}{p_2}$. In cases where the RSD data are incorporated (purple, red, and yellow contours), the $\sigma_{8,0}$ parameter is also displayed.

is conducted with the reciprocal of p_2 , in such a way that the limit effectively becomes $1/p_2 \rightarrow 0^+$.

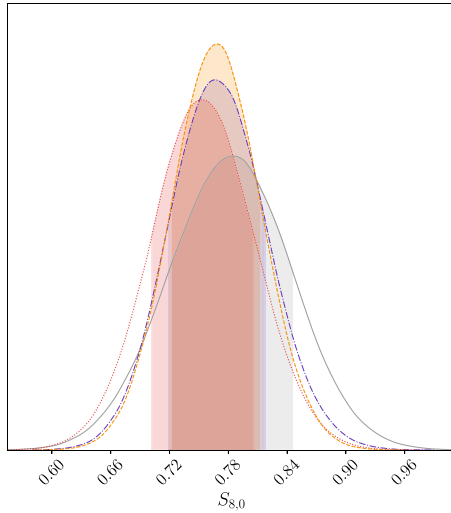
The posterior distributions and confidence levels of the constrained parameters are depicted in Fig. 3. Similar to the f_1 CDM model, the highest value for the H_0 parameter is obtained when combining PN⁺ +RSD data as indicated by the precise values in Table 3. However, in this case, the lowest H_0 value, specifically $H_0 = 66.5^{+2.2}_{-1.3}$ km s⁻¹Mpc⁻¹, is obtained for the CC+BAO+RSD combination. Overall, in this scenario, the parameter values trend slightly lower compared to the f_1 CDM model.

Regarding the $\Omega_{m,0}$ parameter, we observe a similar trend to that of the H_0 parameter, with lower values being reported. However, the consistent pattern persists, where data sets that include RSD values yield lower values for $\Omega_{m,0}$ compared to their counterparts that do not incorporate RSD data. The anticorrelation between the H_0 parameter and the $\Omega_{m,0}$ parameter is still visible. However, in comparison to the f_1 CDM model, the anticorrelation between $\Omega_{m,0}$ and the model parameter is not as pronounced, resulting in a higher degree of degeneracy between these two parameters.

The $1/p_2$ parameter values obtained are slightly higher than those in the f_1 CDM model. Additionally, in contrast to the previous model,

Table 4. Exact $S_{8,0}$ values corresponding to various data sets for the f_2 CDM model.

Data sets	$S_{8,0}$
— RSD	$0.784^{+0.061}_{-0.065}$
- - - CC + BAO + RSD	$0.765^{+0.052}_{-0.046}$
⋯⋯ PN ⁺ & SH0ES + RSD	$0.753^{+0.054}_{-0.051}$
- - - CC + PN ⁺ & SH0ES + BAO + RSD	$0.768^{+0.045}_{-0.046}$

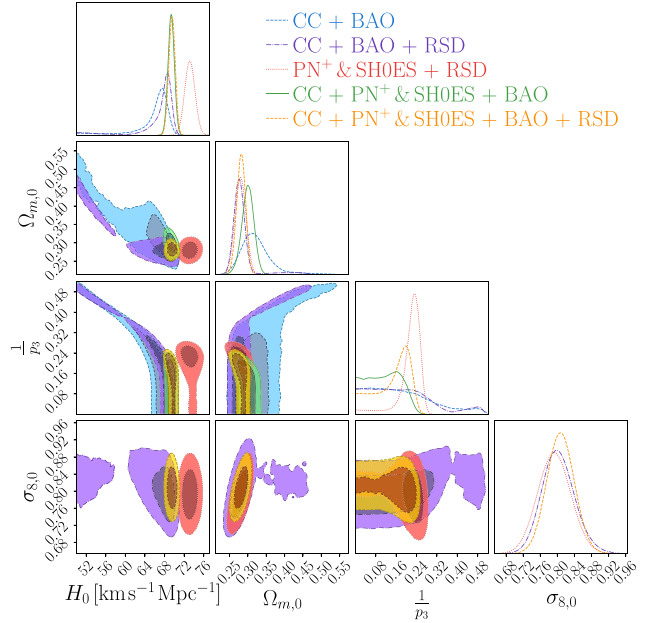

Figure 4. Posterior distribution for the $S_{8,0}$ parameter in the f_2 CDM model. Legend: grey denotes the RSD data, purple corresponds to CC+BAO+RSD data, red represents the PN⁺ & SH0ES + RSD data set, while orange indicates CC + PN⁺ & SH0ES + BAO + RSD data.

the Λ CDM limit does not fall within the 1σ region of the Λ CDM, indicating a slight deviation from the Λ CDM model.

Furthermore, the parameter $\sigma_{8,0}$ continues to exhibit the same trend observed previously, with lower values consistently reported. Notably, we observe a correlation between this parameter and $\Omega_{m,0}$ across all data sets. However, the degeneracy between $\sigma_{8,0}$ and H_0 remains valid in this model. Similar trend to previous model is also seen with regards G_{eff} , where $G_{\text{eff}} \approx G_N$ where

$$G_{\text{eff}} = \frac{G_N}{1 + \frac{1}{2}\alpha_2 p_2 \sqrt{\frac{T_0}{T}} \text{Exp}\left[-p_2 \sqrt{\frac{T}{T_0}}\right]}. \quad (43)$$

In the cases where RSD is included, we once again calculate the quantity $S_{8,0}$, and we obtain slightly different results compared to the previous model as shown in Table 4 and Fig. 4. Similar results are obtained for the CC+BAO+RSD data sets, but we observe a lower value for the PN⁺ & SH0ES+RSD combination. Conversely, we obtain higher values for the RSD data set on its own and for CC + PN⁺ & SH0ES + BAO + RSD, with the maximum value being achieved for the former data set, whereas previously it had exhibited the minimum value. Lastly, we again check the correlation between the parameters p and $S_{8,0}$ where the degeneracy in the RSD data persists, however, when the PN⁺ & SH0ES data are combined with the RSD data, we now notice an anticorrelation that was not previously evident. A visual representation of these findings can be found in Appendix A, specifically in Fig. A1.


Figure 5. Confidence contours and posterior distributions for the f_3 CDM model (exponential model) parameters, including H_0 , $\Omega_{m,0}$, and $\frac{1}{p_3}$. In cases where the RSD data are incorporated (purple, red, and yellow contours), the $\sigma_{8,0}$ parameter is also displayed.

4.3 Exponential model

The next model under consideration in this analysis is the exponential model, hereafter f_3 CDM, which draws inspiration from previous works on $f(\dot{R})$ (Linder 2009). In fact, a variant of the Linder model is proposed in Nesseris et al. (2013), where the square root in the exponential form is no longer present. In this case, \mathcal{F}_3 is expressed as an exponential function with two model constants, α_3 and p_3 , along with the current torsion scalar T_0 and the variable T , such that

$$\mathcal{F}_3 = \alpha_3 T_0 (1 - \text{Exp}[-p_3 T/T_0]). \quad (44)$$

By evaluating the Friedmann equation at the present time, we can determine the value of the constant α_3 , which is calculated as

$$\alpha_3 = \frac{1 - \Omega_{m,0} - \Omega_{r,0}}{(1 + 2p_3)e^{-p_3} - 1}. \quad (45)$$

Therefore, by substituting equation (45) into the modified Friedmann equation (12), we can derive the Friedmann equation for this model, which can be solved numerically

$$E^2(z) = \Omega_{m,0}(1+z)^3 + \Omega_{r,0}(1+z)^4 + \frac{1 - \Omega_{m,0} - \Omega_{r,0}}{(1 + 2p_3)e^{-p_3} - 1} \times [(1 + 2p_3 E^2(z)) \text{Exp}[-p_3 E^2(z)] - 1]. \quad (46)$$

This model exhibits behaviour similar to the Linder model, where the Λ CDM limit is approached as $p_3 \rightarrow \infty$. Therefore, as previously discussed, we perform the analyses using $1/p_3$ to ensure numerical stability, as this approach aligns with the previous model. In this case the G_{eff} is defined as

$$G_{\text{eff}} = \frac{G_N}{1 + \alpha_3 p_3 \text{Exp}\left[-p_3 \frac{T}{T_0}\right]}, \quad (47)$$

where this model exhibits a similar trend to the previous ones.

The confidence levels and the posteriors are found in Fig. 5, whilst the constrained values are found in Table 5. Removing the square root

Table 5. Exact results for f_3 model that include the parameters H_0 , $\Omega_{m,0}$, and $\frac{1}{p_3}$. The $\sigma_{8,0}$ parameter and the nuisance parameter M , are provided for data sets that include RSD or PN⁺ & SH0ES, respectively otherwise, they are left empty.

Data Sets	H_0 [km s ⁻¹ Mpc ⁻¹]	$\Omega_{m,0}$	$\frac{1}{p_3}$	$\sigma_{8,0}$	M
CC + BAO	$67.5^{+1.7}_{-2.3}$	$0.311^{+0.039}_{-0.034}$	$0.058^{+0.182}_{-0.056}$	–	–
CC + BAO + RSD	$68.6^{+1.3}_{-1.9}$	$0.276^{+0.016}_{-0.015}$	$0.026^{+0.214}_{-0.025}$	$0.798^{+0.040}_{-0.036}$	–
PN ⁺ & SH0ES + RSD	$73.2^{+1.0}_{-1.1}$	$0.280^{+0.014}_{-0.015}$	$0.232^{+0.027}_{-0.031}$	$0.793^{+0.038}_{-0.039}$	-19.25 ± 0.11
CC + PN ⁺ & SH0ES + BAO	$69.34^{+0.65}_{-0.64}$	$0.300^{+0.017}_{-0.016}$	$0.160^{+0.029}_{-0.126}$	–	$-19.34^{+0.24}_{-0.31}$
CC + PN ⁺ & SH0ES + BAO + RSD	$69.54^{+0.64}_{-0.66}$	0.282 ± 0.012	$0.197^{+0.038}_{-0.092}$	0.807 ± 0.032	$-19.38^{+0.20}_{-0.19}$

Table 6. Exact $S_{8,0}$ values corresponding to various data sets for the f_3 CDM model.

Data sets	$S_{8,0}$
— RSD	$0.744^{+0.079}_{-0.066}$
- - - CC + BAO + RSD	$0.762^{+0.059}_{-0.052}$
⋯⋯ PN ⁺ & SH0ES + RSD	$0.765^{+0.053}_{-0.052}$
- - - CC + PN ⁺ & SH0ES + BAO + RSD	0.782 ± 0.043

component has notably influenced the constraints, particularly on $\sigma_{8,0}$. The $\sigma_{8,0}$ values are considerably higher in this model compared to those reported in f_2 CDM.

As with f_1 CDM, the model f_3 CDM exhibits the highest constrained value for H_0 when considering the PN⁺ & SH0ES+RSD data ($H_0 = 73.2^{+1.1}_{-1.14}$ km s⁻¹Mpc⁻¹) combination, while the lowest value is obtained for CC+BAO data ($H_0 = 67.5^{+1.7}_{-2.3}$ km s⁻¹Mpc⁻¹). Similarly, the density parameter $\Omega_{m,0}$ also shows the highest value for CC+BAO ($\Omega_{m,0} = 0.311^{+0.039}_{-0.034}$) and the lowest for CC+PN⁺ & SH0ES+BAO ($\Omega_{m,0} = 0.211 \pm 0.012$). In addition, the RSD data appear to impose more stringent constraints on these parameters, particularly on the density parameter.

Regarding the model parameter, unlike in the case of f_2 CDM, we observe that its range within the 1σ and 2σ confidence intervals is narrower. In this instance, the 2σ interval spans from 0 to a maximum of 0.5, but akin in f_2 CDM still lie within 2σ of the Λ CDM limit.

However, the most significant difference in this model becomes evident in the $\sigma_{8,0}$ parameter. This model reports a higher value for this parameter. Furthermore, a correlation between the parameters $\Omega_{8,0}$ and $\sigma_{8,0}$ is now evident, which was not observed in the f_1 CDM model but is apparent in the f_2 CDM model. The degeneracy between the H_0 and $\sigma_{8,0}$ parameters, however, remains apparent, as observed in the other cases.

The highest value for $\sigma_{8,0}$ is obtained for the CC + PN⁺ & SH0ES + BAO + RSD combination, yielding a value of $\sigma_{8,0} = 0.793 \pm 0.035$, as detailed in Table 6 and illustrated in Fig. 6. Conversely, the lowest value for $S_{8,0}$ is achieved when using the PN⁺ & SH0ES+RSD data, as indicated in both tables presenting the $\sigma_{8,0}$ and $S_{8,0}$ values respectively. Additionally, consistent with previous models, the value for $S_{8,0}$ obtained using the RSD data alone is relatively low, akin to the f_1 CDM model.

The final piece of analysis for the f_3 CDM model is depicted in Fig. A1 in Appendix A. Here, we observe a degenerate relationship between the parameters p_3 and $S_{8,0}$, which transitions into an anticorrelation at higher values of p_3 . This effect is particularly noticeable for the data sets that incorporate BAO measurements.

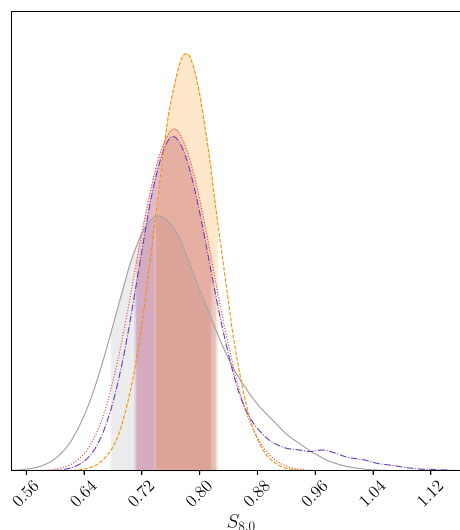


Figure 6. Posterior distribution for the $S_{8,0}$ parameter in the f_3 CDM model. Legend: grey denotes the RSD data, purple corresponds to CC+BAO+RSD data, red represents the PN⁺ & SH0ES + RSD data set, while orange indicates CC + PN⁺ & SH0ES + BAO + RSD data.

4.4 Logarithmic model

The last model, hereafter referred to as f_4 CDM, which was proposed by Bamba et al. (2011), is known as the logarithmic model. This model is characterized by the following expression

$$\mathcal{F}_4(T) = \alpha_4 T_0 \sqrt{\frac{T}{p_4 T_0}} \log \left[\frac{p_4 T_0}{T} \right], \quad (48)$$

where α_4 and p_4 are the two model constants. To determine the constant α_4 , we follow the standard procedure of evaluating the Friedmann equation (13), at the current time, resulting in

$$\alpha_4 = -\frac{(1 - \Omega_{m,0} - \Omega_{r,0}) \sqrt{p_4}}{2}. \quad (49)$$

This simplifies the Friedmann equation to a more straightforward form

$$E^2(z) = \Omega_{m_0} (1+z)^3 + \Omega_{r_0} (1+z)^4 + (1 - \Omega_{m,0} - \Omega_{r_0}) E(z) \quad (50)$$

Interestingly, the additional model parameter p_4 does not feature in equation (50), indicating that at background level this parameter cannot be constrained. None the less, it is noteworthy that this parameter does come into play at the perturbative level, particularly in the context of linear matter perturbations, as illustrated in equation (22), where p_4 is featured in the G_{eff} term, in particular in the \mathcal{F}_T term. Therefore, as can be seen in the contour plots in Fig. 7 and Table 7,

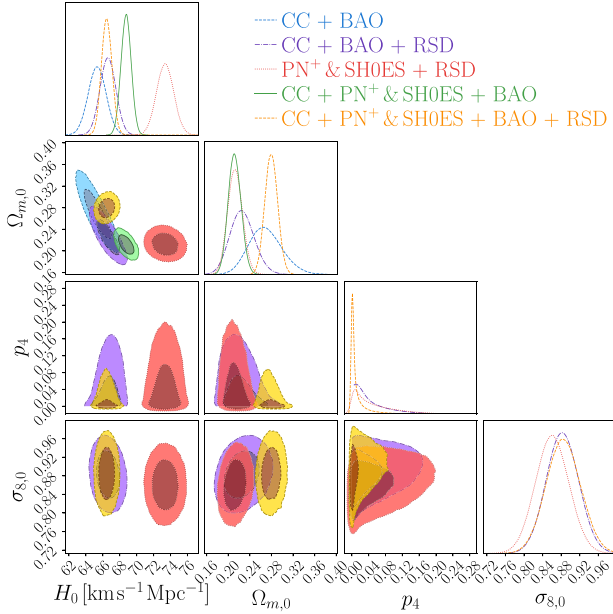


Figure 7. Confidence contours and posterior distributions for the f_4 CDM model (logarithmic model) parameters, including H_0 and $\Omega_{m,0}$. In cases where the RSD data are incorporated (purple, red, and yellow contours), the $\sigma_{8,0}$ and p_4 parameter is also displayed.

the p_4 parameter does not feature in the CC+BAO (blue contours) and CC+PN⁺ & SHOES+BAO (green contour), but is relevant for the rest of the data sets that include the RSD data.

The inclusion of the logarithmic function makes a noticeable difference in this model, particularly when examining the relationship between the H_0 and $\Omega_{m,0}$ parameters. In this case, the relationship exhibits a distinct anticorrelation, which is more pronounced in this model compared to the previous ones. Interestingly, this results in more extreme values for the H_0 parameter, with the highest value achieved being $H_0 = 73.40^{+0.97}_{-1.06} \text{ km s}^{-1} \text{ Mpc}^{-1}$ for the PN⁺ & SHOES data set. Conversely, the lowest value was obtained for the CC+BAO data set, with a value of $H_0 = 65.3^{+1.1}_{-1.0} \text{ km s}^{-1} \text{ Mpc}^{-1}$.

The values of $\Omega_{m,0}$ also exhibit significant variations. In this case, the lowest value is achieved for the CC+PN⁺ & SHOES+BAO data set, while the highest is for the same data set but with the inclusion of the RSD data (i.e. CC+PN⁺ & SHOES+BAO+RSD). It seems like when the RSD model is included in the data sets, either the H_0 parameter or the $\Omega_{m,0}$ is driven to higher values when compared to its counterpart data set without the RSD data set. Therefore, in this model, it appears that the presence of RSD data influences the understanding of the Universe’s current acceleration and the role of baryonic and dark matter components within this model. This influence can be attributed to RSD data’s exceptional sensitivity to

the distribution of matter, which in turn plays a crucial role in shaping the large-scale structure of the Universe through its gravitational interactions.

A particularly intriguing aspect of this model, which sets it apart from the $f_1 - f_3$ CDM models, is the absence of a Λ CDM limit both at the background and perturbative levels. At the perturbative level, the parameter p_4 appears in G_{eff} , which is expressed as

$$G_{\text{eff}} = \frac{G_N}{1 + \frac{\alpha_4}{2\sqrt{p_4}} \left(\frac{T}{T_0}\right)^{-\frac{1}{2}} \ln\left(\frac{p_4 T_0}{T}\right) + \alpha_4 \sqrt{\frac{T_0}{p_4 T}}}. \quad (51)$$

In this scenario, no choice of the value of p_4 can reproduce a Λ CDM model. Furthermore, given that it appears in the logarithmic function, the parameter p_4 must be greater than zero, as illustrated in Fig. 7. However, Table 7, still shows that the value of p_4 is very close to zero, which makes G_{eff} , slightly higher than G_N , throughout the redshift span.

The $\sigma_{8,0}$ parameter in this model appears to be significantly higher when compared to the other models. Interestingly, the CC+BAO+RSD and CC+PN⁺ & SHOES+BAO+RSD data sets both report the same value for $\sigma_{8,0}$, suggesting that the inclusion of PN⁺ & SHOES data has a minimal effect on this parameter in this particular model. The CC+PN⁺ & SHOES + RSD data set reports a slightly lower value for $\sigma_{8,0}$. This observation leads us to consider Fig. 8 and Table 8, where we can see the values obtained for this parameter. In this case, it’s evident that $\Omega_{m,0}$ has a notable impact on this parameter. While $\sigma_{8,0}$ is the same for CC+BAO+RSD and CC+PN⁺ & SHOES+BAO+RSD, $\sigma_{8,0}$ varies significantly between the two data sets. A higher value is obtained when PN⁺ & SHOES data are included, making it the data set with the highest $\sigma_{8,0}$ value.

The last Fig. A1 (Appendix A), shows the relationship between the $\sigma_{8,0}$ and p_4 . As expected based on previous models, we observe a degeneracy between these parameters, particularly concentrated within the 1σ region near $p_4 = 0$.

5 ANALYSIS

To evaluate how well each f_i CDM model performs with various data sets, we employ several statistical measures. First, we calculate the minimum χ^2_{min} values for each model and data set. These values are derived from the maximum likelihood L_{max} , with the relationship being $\chi^2_{\text{min}} = -2 \ln L_{\text{max}}$. A lower χ^2_{min} indicates a better fit of the model to the data.

In addition to χ^2_{min} , we also compare the models to the standard Λ CDM model using two criteria: the Akaike Information Criterion (AIC) and the Bayesian Information Criterion (BIC). These criteria consider both the model’s goodness of fit, represented by χ^2_{min} , and its complexity, which is determined by the number of parameters n . The AIC is calculated as

$$\text{AIC} = \chi^2_{\text{min}} + 2n, \quad (52)$$

Table 7. Exact results for f_4 model that include the parameters H_0 and $\Omega_{m,0}$. The $\sigma_{8,0}$ and p_4 parameter together with the nuisance parameter M , are provided for data sets that include RSD or PN⁺ & SHOES, respectively otherwise, they are left empty.

Data Sets	H_0 [km s ⁻¹ Mpc ⁻¹]	$\Omega_{m,0}$	p_4	$\sigma_{8,0}$	M
CC + BAO	$65.3^{+1.1}_{-1.0}$	$0.266^{+0.031}_{-0.030}$	–	–	–
CC + BAO + RSD	$66.62^{+0.87}_{-0.91}$	$0.224^{+0.022}_{-0.020}$	$0.005^{+0.031}_{-0.000}$	0.883 ± 0.033	–
PN ⁺ & SHOES + RSD	$73.40^{+0.97}_{-1.06}$	$0.213^{+0.013}_{-0.014}$	$0.007^{+0.057}_{-0.013}$	$0.860^{+0.036}_{-0.035}$	$-19.09^{+0.16}_{-0.23}$
CC + PN ⁺ & SHOES + BAO	$68.78^{+0.58}_{-0.63}$	0.211 ± 0.012	–	–	-19.388 ± 0.017
CC + PN ⁺ & SHOES + BAO + RSD	$66.41^{+0.67}_{-0.56}$	$0.280^{+0.011}_{-0.012}$	$0.0015^{+0.009}_{-0.005}$	$0.883^{+0.038}_{-0.036}$	$-19.57^{+0.16}_{-0.15}$

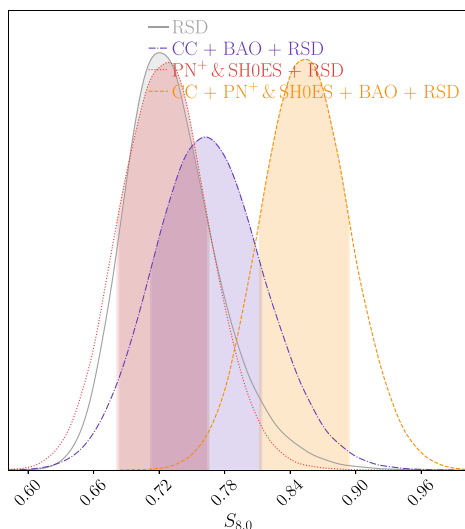


Figure 8. Posterior distribution for the $S_{8,0}$ parameter in the f_4 CDM model. Legend: grey denotes the RSD data, purple corresponds to CC+BAO+RSD data, red represents the PN⁺ & SH0ES + RSD data set, while orange indicates CC + PN⁺ & SH0ES + BAO + RSD data.

Table 8. Exact $S_{8,0}$ values corresponding to various data sets for the f_4 CDM model.

Data sets	$S_{8,0}$
— RSD	$0.720^{+0.047}_{-0.036}$
--- CC + BAO + RSD	$0.762^{+0.052}_{-0.050}$
..... PN ⁺ & SH0ES + RSD	$0.730^{+0.035}_{-0.048}$
-- CC + PN ⁺ & SH0ES + BAO + RSD	$0.854^{+0.040}_{-0.042}$

where a lower AIC value indicates a model that fits the data better while accounting for complexity. It penalizes models with more parameters, even if they exhibit a superior data fit. On the other

Table 9. Comparison of χ^2_{\min} and differences in AIC and BIC between the models and Λ CDM (i.e Δ AIC and Δ BIC). On the left-hand side, results are presented for CC+BAO, while the right-hand side includes RSD.

	χ^2_{\min}	CC+BAO Δ AIC	Δ BIC	χ^2_{\min}	CC+BAO+RSD Δ AIC	Δ BIC
Λ CDM	20.93	0	0	37.14	0	0
f_1 CDM	20.87	1.94	1.61	37.04	1.91	3.15
f_2 CDM	20.93	2.00	1.66	35.41	0.28	1.52
f_3 CDM	20.93	2.00	1.66	37.21	2.08	3.32
f_4 CDM	27.95	7.02	7.02	42.03	6.89	8.14

Table 10. Comparison of χ^2_{\min} and differences in AIC and BIC between the models and Λ CDM (i.e Δ AIC and Δ BIC). On the left-hand side, results are presented for PN⁺ & SH0ES+RSD, whilst in the middle CC+PN⁺ & SH0ES+BAO. On the right-hand side, the results for CC+PN⁺ & SH0ES+BAO+RSD are displayed.

	PN ⁺ & SH0ES+RSD			CC+PN ⁺ & SH0ES+BAO			CC+PN ⁺ & SH0ES+BAO+RSD		
	χ^2_{\min}	Δ AIC	Δ BIC	χ^2_{\min}	Δ AIC	Δ BIC	χ^2_{\min}	Δ AIC	Δ BIC
Λ CDM	1550.20	0	0	1572.60	0	0	1590.71	0	0
f_1 CDM	1549.52	1.32	2.56	1572.56	1.96	1.78	1590.56	1.85	3.10
f_2 CDM	1541.46	-6.74	-5.50	1572.50	1.89	1.72	1587.49	-1.21	0.04
f_3 CDM	1543.13	-5.07	-3.84	1572.31	1.71	1.53	1588.33	-0.37	0.88
f_4 CDM	1539.71	-8.48	-7.25	1586.67	14.07	14.07	1663.50	74.80	76.04

hand, the BIC is given by

$$\text{BIC} = \chi^2_{\min} + n \ln m, \quad (53)$$

where m is the sample size of the observational data combination. Similar to the AIC, the BIC aims to balance data fit against model complexity. However, it imposes a heavier penalty on models with more parameters as the sample size increases. Therefore, by comparing the AIC and BIC values of different models, we can determine which model is better supported by the data. Generally, models with lower AIC and BIC values are preferred, provided the differences are significant.

To assess the performance of different models using various combinations of data sets, we calculate the differences in AIC and BIC between each model (f_i CDM) and the reference model, which is the Λ CDM model. This comparison helps us understand how well each model aligns with the standard model of cosmology. The differences in AIC and BIC are denoted as $\Delta\text{AIC} = \Delta\chi^2_{\min} + 2\Delta n$, and $\Delta\text{BIC} = \Delta\chi^2_{\min} + \Delta n \ln m$, respectively. These metrics quantify how each model deviates from the reference model (in this case, Λ CDM) in which, smaller values of ΔAIC and ΔBIC indicate that a model, along with its chosen data set, is more similar to the Λ CDM model, suggesting better performance. Table 9 provides a comparison of these metrics for two specific data set combinations: CC+BAO and CC+BAO+RSD. Similarly, Table 10 allows us to compare the AIC and BIC values for three different sets of data combinations: PN⁺&SH0ES+RSD, CC+PN⁺&SH0ES+BAO and CC+PN⁺&SH0ES+BAO+RSD.

Table 9 highlights that the CC+BAO data set generally exhibits a lower χ^2_{\min} , implying a better fit without RSD. However, closer inspection of the ΔAIC values for CC+BAO+RSD reveals a significant reduction, indicating the increased favourability of this data set in certain cases. It is noteworthy that the BIC values for CC+BAO+RSD are slightly higher due to more data points. Regarding model performance, the f_2 model consistently presents the lowest ΔAIC and ΔBIC values, signifying its reliable performance. The f_1 and f_3 models also perform well, showing suitability for this data set. In contrast, the f_4 CDM model tends to lean towards the

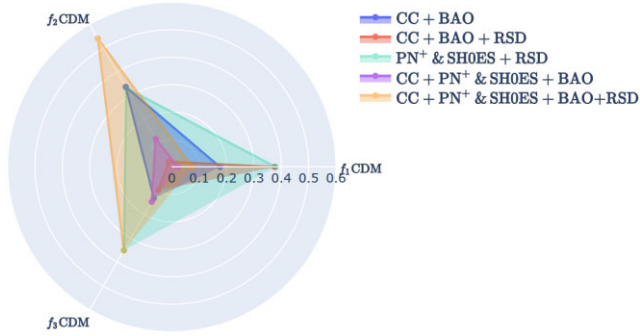


Figure 9. Distances measured in standard deviations (σ units) between the constrained H_0 values obtained from the $f_1 - 3$ CDM models and their corresponding values in the Λ CDM model.

Λ CDM model, as suggested by higher Δ AIC and Δ BIC values, indicating weaker data support for f_4 CDM.

Table 10 reveals intriguing findings. Starting with the $\text{PN}^+ \& \text{SH0ES} + \text{RSD}$ data set, the models are favoured over Λ CDM, as reflected in negative AIC and BIC values, indicating a preference for the models. While this preference is not consistent across all models for the $\text{CC} + \text{PN}^+ \& \text{SH0ES} + \text{BAO}$ data set, it re-emerges for $\text{CC} + \text{PN}^+ \& \text{SH0ES} + \text{BAO} + \text{RSD}$, especially pronounced for the f_2 and f_3 models. This suggests that, for these two models within the $\text{CC} + \text{PN}^+ \& \text{SH0ES} + \text{BAO} + \text{RSD}$ data sets, the data lean towards favouring these models over Λ CDM. However, it is important to note that the evidence is not strong enough to definitively favour these models over Λ CDM. The BIC values, although not negative, are nearly zero, indicating an inconclusive preference for either model. Conversely, the f_4 model shows strong disfavour and preference for Λ CDM, prompting its exclusion from further analysis, shifting our focus to the first three models.

The previous analysis is further supported by Fig. 9. In this figure, we compare the values of H_0 obtained from $f_i(T)$ ($i = 1, 2, 3$) models to the values of H_0 obtained from the Λ CDM model for the same data sets, as shown in Appendix B. This visualization illustrates the variations in H_0 across different data sets, expressed in terms of σ units, with each data set represented by a distinct colour. For each $f(T)$ model, we observe that the H_0 values fall within 1σ of the corresponding Λ CDM values. Therefore, the H_0 values obtained for the different data sets are consistent with those of Λ CDM.

The discrepancy between the locally measured expansion rate of the Universe and the values inferred from observations of the CMB has prompted us to extend our previous analysis. We aim to investigate how these models perform with different values of H_0 , specifically considering the P18 value $H_0^{\text{P18}} = 67.4 \pm 0.5 \text{ km s}^{-1} \text{ Mpc}^{-1}$ (Aghanim et al. 2020) and R22 value of $H_0^{\text{R22}} = 73.30 \pm 1.4 \text{ km s}^{-1} \text{ Mpc}^{-1}$ (Riess et al. 2022), as shown in Fig. 10 (in both cases we include the Λ CDM values from Appendix B).

In Fig. 10, on the left-hand-side, we observe that the H_0 values are within 3.5σ of the P18 value, except for the $\text{PN}^+ \& \text{SH0ES}$ data set. However, in the figure, we also include values from RSD data only. This high σ value for the $\text{PN}^+ \& \text{SH0ES}$ data is expected, as it dominates and is included within the R22 value itself. Similarly, in Fig. 10, on the right-hand side, we also see that the H_0 values are within approximately 3.5σ of the R22 value. This suggests that the H_0 values obtained from our analysis fall approximately midway between the two quoted values of P18 and R22.

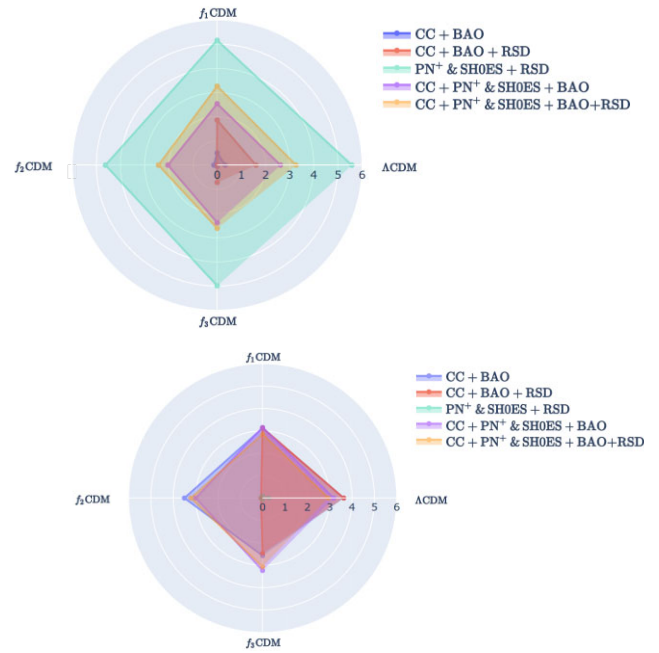


Figure 10. Distances measured in standard deviations (σ units) between the constrained H_0 values obtained from the $f_1 - 3$ CDM models together with Λ CDM and the H_0^{P18} value on the left-hand side and H_0^{R22} on the right-hand side.

Additionally, we present a comprehensive comparison of our $S_{8,0}$ values with the Planck Collaboration's $S_{8,0}^{\text{P18}} = 0.834 \pm 0.016$ (Aghanim et al. 2020) and Kilo-Degree Survey (KiDS-1000) value $S_{8,0}^{\text{KiDS}} = 0.766^{+0.020}_{-0.014}$ (Asgari et al. 2021). The results are summarized in Fig. 11, highlighting the impact of different data sets on our models. a comparison between $S_{8,0}$ values in our models and those in Λ CDM is presented in Fig. 12. Furthermore, Fig. 13 provides a comparison of $S_{8,0}^{\text{P18}}$ and $S_{8,0}^{\text{KiDS}}$ with values derived from $f_1 - 3$ CDM models in σ units. When compared to $S_{8,0}^{\text{P18}}$, on the left-hand side of same figure, we find a maximum discrepancy of 2σ between the values derived from Λ CDM and the $f(T)$ models, and this is primarily observed when considering the RSD data set in isolation. However, as more data sets are included, particularly $\text{CC} + \text{PN}^+ \& \text{SH0ES} + \text{BAO} + \text{RSD}$, which is the most comprehensive data set in our analysis, this discrepancy diminishes.

On the right-hand panel of Fig. 13, where we compare with $S_{8,0}^{\text{KiDS}}$, the discrepancy is further reduced to 1σ , indicating that the values obtained from our analysis are closer to the KiDS-1000 measurements than the Planck values. Notably, $\text{CC} + \text{PN}^+ \& \text{SH0ES} + \text{BAO} + \text{RSD}$ performs well in this regard, significantly reducing the σ values. Consequently, it appears that with the combination of the available data sets and the $f_1 - 3(T)$ models, the tensions between the CMB and locally determined values are reduced, especially for $S_{8,0}$.

6 CONCLUSION

The most popular models in the literature of TG have been probed in this work against the latest expansion data as well as RSD measurements which has expanded the constraint profile of each of these models, as well as their impact on physically observable cosmological parameters. For the local data being used, we had a combination of CC, $\text{PN}^+ \& \text{SH0ES}$, and BAO data. Besides being

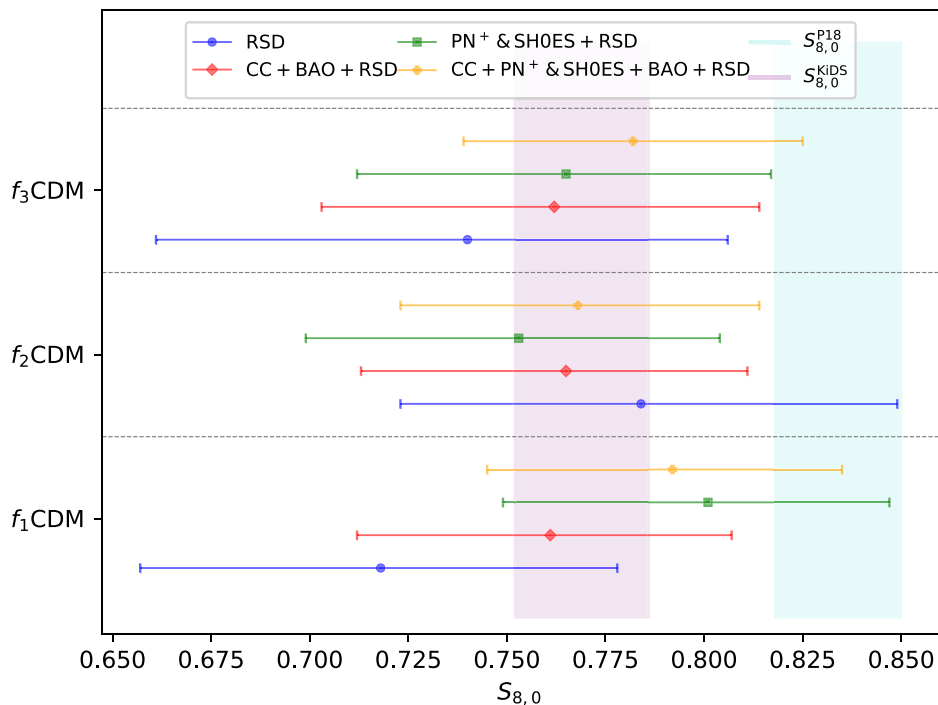


Figure 11. A whisker plot illustrating the constrained values of $S_{8,0}$ as derived from the $f_1 - 3$ CDM models. The cyan and purple vertical bands depict the respective 1σ ranges of $S_{8,0}^{P18}$ and $S_{8,0}^{KIDS}$, whilst the error bars illustrate the inferred model dependent 1σ constraints from each respective data set.

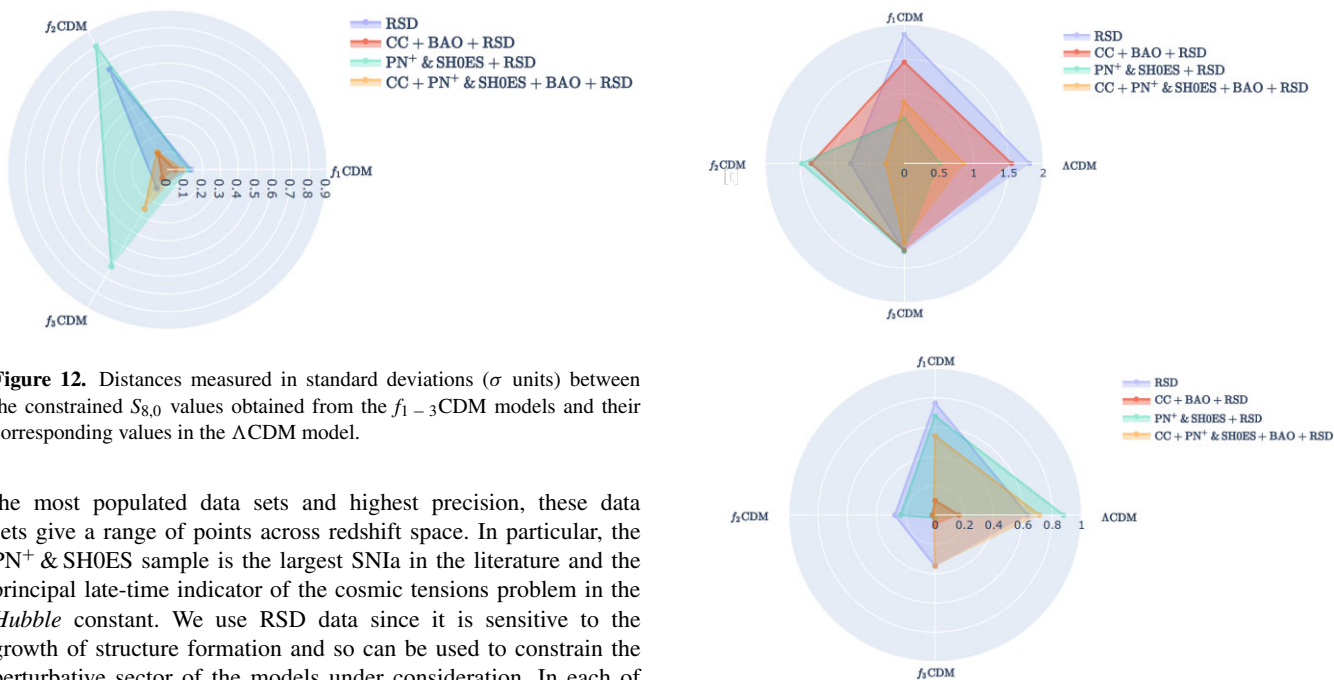


Figure 12. Distances measured in standard deviations (σ units) between the constrained $S_{8,0}$ values obtained from the $f_1 - 3$ CDM models and their corresponding values in the Λ CDM model.

the most populated data sets and highest precision, these data sets give a range of points across redshift space. In particular, the PN⁺ & SH0ES sample is the largest SNIa in the literature and the principal late-time indicator of the cosmic tensions problem in the *Hubble* constant. We use RSD data since it is sensitive to the growth of structure formation and so can be used to constrain the perturbative sector of the models under consideration. In each of these models and for each data set combination, we performed a full MCMC analysis obtaining constraints on all the cosmological parameters. Additionally, we compared the performance of each model for each data set combination against the standard Λ CDM model using the χ^2_{\min} , AIC, and BIC statistical indicators. Given the increasing tension being reported by various studies on the value of the *Hubble* constant, we study its constrained value, but also consider the effect of a changing *Hubble* constant will have on the parameters related to the growth of structure formation, namely, $\sigma_{8,0}$ and $S_{8,0}$.

Figure 13. Distances measured in standard deviations (σ units) between the constrained $S_{8,0}$ values obtained from the $f_1 - 3$ CDM models together with Λ CDM and the $S_{8,0}^{P18}$ value on the left-hand side and $S_{8,0}^{R22}$ on the right-hand side.

For reference purposes, we provide the Λ CDM constraint values for each of the data set combinations we consider. This is done in Appendix B, which is important for estimating the statistical indicators in the model sections. Generally, our analysis shows consistency

with Λ CDM but some differences do arise, which may show further distinction as the precision in measurements improves. As one might expect, the strongest data set in the analysis was the PN⁺ & SHOES sample since it consistently drastically reduces the statistical error for any baseline data set. On the other hand, the statistical metrics show that expansion data CC+PN⁺ & SHOES+BAO combination provides evidence for Λ CDM while when the RSD data set is included, there is marginal preference for the $f(T)$ model under consideration. Moreover, they offer best fits on cosmic parameters that are more aligned with a higher *Hubble* constant. As for the $S_{8,0}$ parameter, the models are largely consistent with the most recent reported values in the literature.

This analysis offers precision insights into the behaviour of these models, but also on TG more generally, when using combinations of this expansion profile and RSD data sets. The study suggests that these models may offer some promise as toy models for modified cosmological scenario model building. We intend to use these observations to extend this analysis to include CMB power spectra as well as other early Universe data sets to more fully assess the competitiveness of these models.

DATA AVAILABILITY

The data used in this study are described in detail in Section 3. For further information on the data underlying this article, contact the corresponding author who will provide any additional information.

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APPENDIX A: MODEL PARAMETER p_i VERSUS $S_{8,0}$ PLOTS

In this section, we present the posteriors together with their confidence regions of the p_i and $S_{8,0}$ parameters to investigate the correlation between the two.

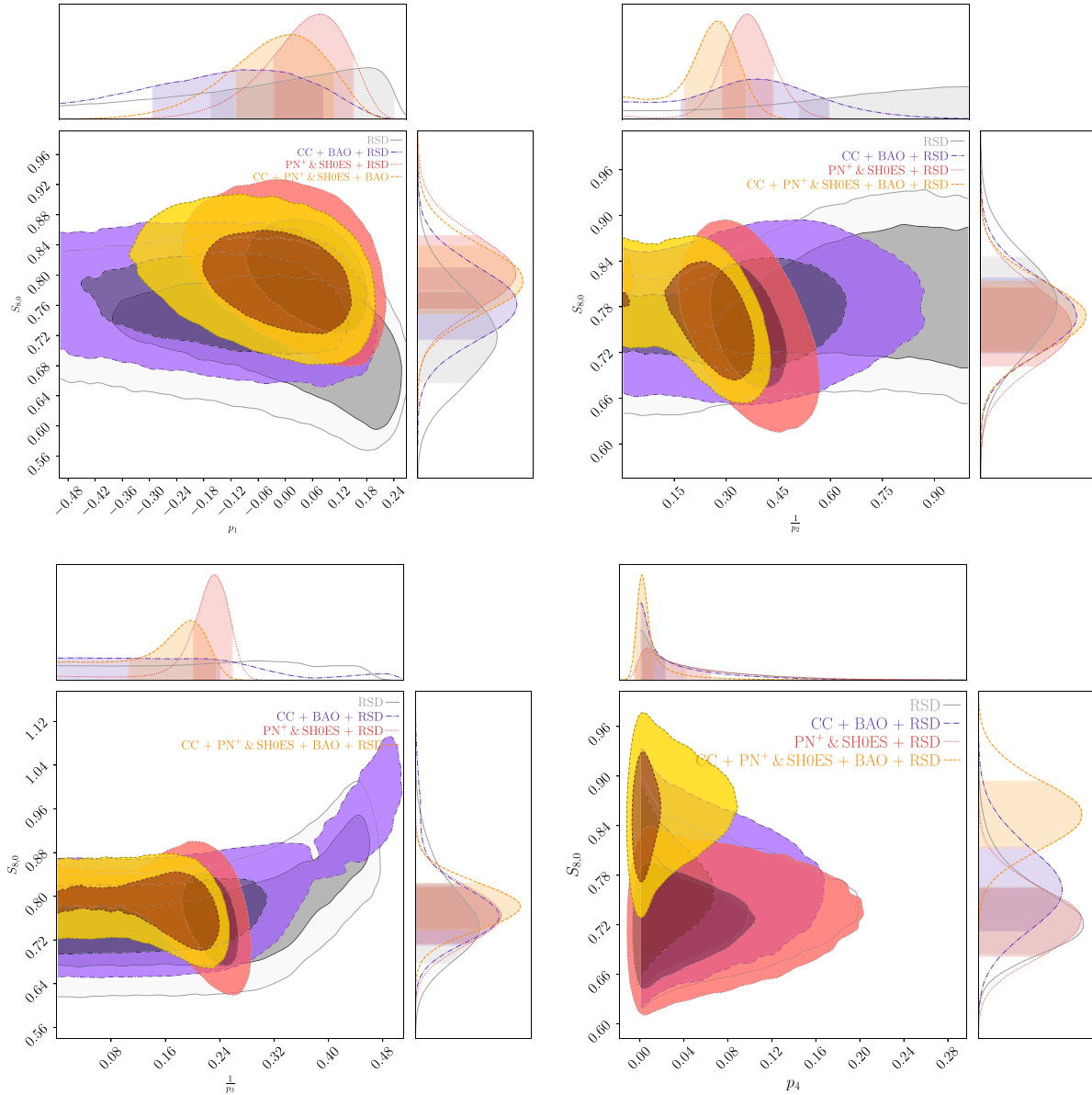


Figure A1. Top left: contour plots showing the relationship between the p_1 parameter and the $S_{8,0}$ parameter for the f_1 CDM model (power-law model). Top right: contour plots showing the relationship between the $\frac{1}{p_2}$ parameter and the $S_{8,0}$ parameter for the f_2 CDM model (Linder model). Left bottom: contour plots showing the relationship between the $\frac{1}{p_3}$ parameter and the $S_{8,0}$ parameter for the f_1 CDM model (exponential model). Right bottom: contour plots showing the relationship between the p_4 parameter and the $S_{8,0}$ parameter for the f_1 CDM model (logarithmic model).

APPENDIX B: Λ CDM MODEL

In Section 5, we provide comparisons between the models and their corresponding Λ CDM values. Here, we present the posterior distributions and confidence regions of the Λ CDM model

in Fig. B1 and provide additional details and precise values in Table B1. In Fig. B2 the posteriors of the $S_{8,0}$ parameter for the Λ CDM is shown, whilst in Table B2 the exact values for the corresponding data sets are presented.

Table B1. Exact results for Λ CDM model that include the parameters H_0 and $\Omega_{m,0}$. The $\sigma_{8,0}$ parameter and the nuisance parameter M , are provided for data sets that include RSD or PN⁺ & SH0ES, respectively otherwise, they are left empty.

Data sets	H_0 [km s ⁻¹ Mpc ⁻¹]	$\Omega_{m,0}$	$\sigma_{8,0}$	M
CC + BAO	67.8 ± 1.1	$0.308^{+0.032}_{-0.029}$	–	–
CC + BAO + RSD	$68.77^{+0.71}_{-0.67}$	$0.276^{+0.013}_{-0.014}$	$0.789^{+0.035}_{-0.033}$	–
PN ⁺ & SH0ES + RSD	$73.71^{+0.97}_{-1.06}$	$0.298^{+0.011}_{-0.012}$	$0.814^{+0.037}_{-0.034}$	$-19.249^{+0.027}_{-0.032}$
CC + PN ⁺ & SH0ES + BAO	$69.47^{+0.59}_{-0.63}$	$0.304^{+0.015}_{-0.014}$	–	-19.375 ± 0.017
CC + PN ⁺ & SH0ES + BAO + RSD	$69.84^{+0.55}_{-0.56}$	$0.288^{+0.010}_{-0.010}$	$0.815^{+0.030}_{-0.033}$	$-19.369^{+0.016}_{-0.017}$

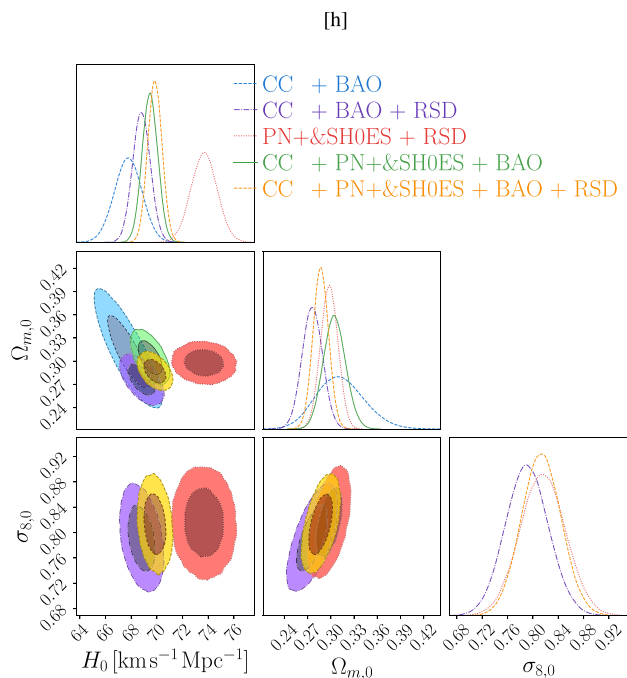
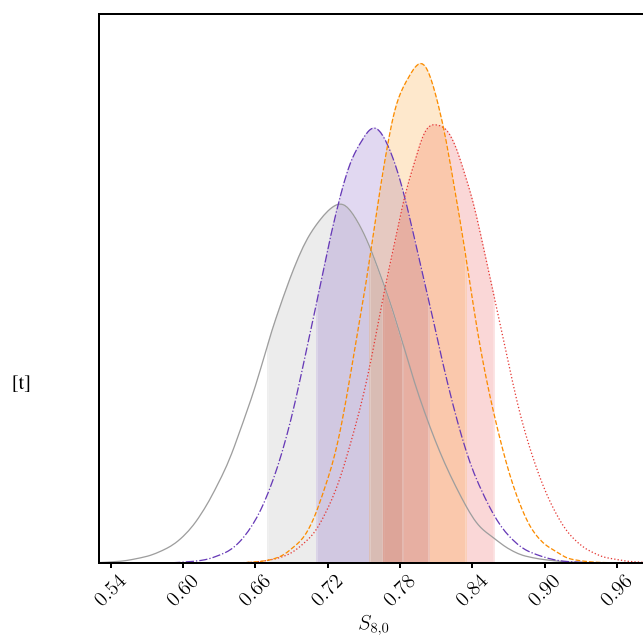

Figure B1. Confidence contours and posterior distributions for the Λ CDM model (logarithmic model) parameters, including H_0 and $\Omega_{m,0}$. In cases where the RSD data are incorporated (purple, red, and yellow contours), the $\sigma_{8,0}$ parameter is also displayed.

Figure B2. Posterior distribution for the $S_{8,0}$ parameter in the Λ CDM model. Legend: grey denotes the RSD data, purple corresponds to CC+BAO+RSD data, red represents the PN⁺ & SH0ES + RSD data set, while orange indicates CC + PN⁺ & SH0ES + BAO + RSD data.

Table B2. Exact $S_{8,0}$ values corresponding to various data sets for the Λ CDM model.

Model	$S_{8,0}$
— RSD	$0.729^{+0.053}_{-0.059}$
- - - CC + BAO + RSD	$0.758^{+0.046}_{-0.047}$
..... PN ⁺ & SH0ES + RSD	$0.809^{+0.050}_{-0.042}$
- - - CC + PN ⁺ & SH0ES + BAO + RSD	$0.797^{+0.038}_{-0.042}$

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