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ORIGINAL ARTICLE

Reappraisal of prognostic factors used in the European Pediatric Soft Tissue Sarcoma Study Group RMS 2005 study for localized rhabdomyosarcoma to optimize risk stratification and generate a prognostic nomogram

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Abstract

Background: The objective of this study was to investigate the role of clinical factors together with *FOXO1* fusion status in patients with nonmetastatic rhabdomyosarcoma (RMS) to develop a predictive model for event‐free survival and provide a rationale for risk stratification in future trials.

Methods: The authors used data from patients enrolled in the European *Pediatric* Soft Tissue Sarcoma Study Group (E*p*SSG) RMS 2005 study (E*p*SSG RMS 2005; EudraCT number 2005‐000217‐35). The following baseline variables were considered for the multivariable model: age at diagnosis, sex, histology, primary tumor site, Intergroup Rhabdomyosarcoma Studies group, tumor size, nodal status, and *FOXO1* fusion status. Main effects and significant second‐order interactions of candidate predictors were included in a multiple Cox proportional hazards regression model. A nomogram was generated for predicting 5‐year event‐free survival (EFS) probabilities.

Results: The EFS and overall survival rates at 5 years were 70.9% (95% confidence interval, 68.6%–73.1%) and 81.0% (95% confidence interval, 78.9%–82.8%), respectively. The multivariable model retained five prognostic factors, including age at diagnosis interacting with tumor size, tumor primary site, Intergroup

The first two authors contributed equally to this article.

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Rhabdomyosarcoma Studies clinical group, and *FOXO1* fusion status. Based on each patient's total score in the nomogram, patients were stratified into four groups. The 5‐year EFS rates were 94.1%, 78.4%, 65.2%, and 52.1% in the low‐risk, intermediate-risk, high-risk, and very-high-risk groups, respectively, and the corresponding 5‐year overall survival rates were 97.2%, 91.5%, 74.3%, and 60.8%, respectively.

Conclusions: The results presented here provide the rationale to modify the E*p*SSG stratification, with the most significant change represented by the replacement of histology with fusion status. This classification was adopted in the new international trial launched by the E*p*SSG.

KEYWORDS

FOXO1 protein, human, nomograms, pediatrics, proportional hazards models, rhabdomyosarcoma, survival analysis

INTRODUCTION

The survival of patients with rhabdomyosarcoma (RMS) has improved in the past 30 years because of the application of a multimodality approach that includes chemotherapy with surgery and/or radiotherapy. Clinical trials coordinated by national and international cooperative groups have helped to refine the treatment and to identify the most active multidrug regimens through randomized studies. A major advance has been the capacity to tailor the treatment strategy according to a series of prognostic factors found to be associated with different levels of the risk of treatment failure. $1-3$ The most powerful adverse risk factor for patients with RMS is the presence of metastases at diagnosis. In this group, the outcome tends to be much poorer, with only one third of patients surviving 3 years after diagnosis.^{[4,5](#page-9-0)}

In the absence of metastatic dissemination, the search for prognostic factors is made difficult by the clinical and biologic heterogeneity of RMS: patients vary in age (with two peaks of incidence: those younger than 6 years and adolescents), the primary tumor arises in many different sites across the body, and its disease extent and involvement of nearby organs and lymph nodes show considerable variation, with consequences for accessibility to local therapy. Two main histologic subtypes are distinguished: embryonal RMS (70% of all RMS) and, with a poorer prognosis, alveolar RMS (ARMS; 20%–30% of all RMS), characterized by the presence of *PAX3/7‐ FOXO1* translocations.

All this information was used by the European *Pediatric* Soft Tissue Sarcoma Study Group (E*p*SSG) to elaborate a risk‐ stratification system that has been used in the RMS 2005 protocol for nonmetastatic RMS. The E*p*SSG risk‐stratification system was based on six prognostic factors: histology, postsurgical stage according to Intergroup Rhabdomyosarcoma Studies (IRS) grouping, primary tumor site, nodal involvement, tumor size, and patient age, as reported in Table [1](#page-2-0).

Since the design of the RMS 2005 protocol, the association between *PAX3/7‐FOXO1* translocation and poorer prognosis has been recognized, overruling the impact of histologic classification and leading to the replacement of histology by *FOXO1* fusion status in the current risk‐stratification system used by the Children's Oncology Group (COG)[3](#page-9-0) and the new E*p*SSG Frontline and Relapse Rhabdomyosarcoma (FaR‐RMS) study (ClinicalTrials.gov identifier NCT04625907).

The objective of the current study presented was to investigate the role of clinical factors together with *FOXO1* fusion status in patients with nonmetastatic RMS treated on the RMS 2005 protocol to develop a predictive model for event‐free survival (EFS) and provide a rationale for risk stratification in future trials.

MATERIALS AND METHODS

Patients and treatments

For this study, we used data from 1733 patients with nonmetastatic rhabdomyosarcoma who were enrolled in the E*p*SSG RMS 2005 study (EudraCT number: 2005‐000217‐35) from October 1, 2005, to December 31, 2016. The data cutoff for last follow‐up was November 15, 2022. Only patients who had complete data were eligible for the analyses. Because *FOXO1* status was not always investigated a priori for patients who had favorable histology (botryoid, embryonal, spindle cells/leiomiomatous), it was assumed to be negative, and we performed a *clinical* and fixed imputation considering *FOXO1* fusion status as negative for the 361 patients who had favorable histology RMS without *FOXO1* fusion data.⁶ Patients who had unfavorable histology (ARMS, solid alveolar, and not otherwise specified) without *FOXO1* fusion data ($n = 47$) and those who had missing clinical data $(n = 9$ without a record of nodal involvement and $n = 16$ with missing tumor size) were excluded, yielding 1661 evaluable patients (see

TABLE 1 Risk grouping stratification and therapy in the European Pediatric Sarcoma Study Group Rhabdomyosarcoma 2005 study.

Risk group	Subgroups	Pathology	Postsurgical stage (IRS group	Site	Node stage	Size and age	Chemotherapy	Delayed surgery	Radiation therapy
Low risk	A	Favorable	I	Any	N ₀	Favorable	VA \times 8 cycles	Not necessary	No
Standard risk	B	Favorable	\mathbf{I}	Any	N ₀	Unfavorable	IVA \times 4 cycles $+ VA \times$ 5 cycles	Not necessary	No
	C	Favorable	II, III	Favorable	N ₀	Any	IVA \times 9 cycles or IVA \times 5 cycles $+ VA \times 4$ cycles if radiotherapy	Yes, if not mutilating	Optional
	D	Favorable	II, III	Unfavorable	N ₀	Favorable	IVA \times 9 cycles	Yes, if not mutilating	Yes
High risk	Ε	Favorable	II, III	Unfavorable	N ₀	Unfavorable	IVA \times 9 cycles vs. IVADo \times 4 cycles $+$ IVA \times 5 cycles with or without maintenance \times 6 cycles	Yes	Yes
	F	Favorable	II, III	Any	N ₁	Any			
	G	Unfavorable	I, II, III	Any	N ₀	Any			
Very high risk	H	Unfavorable	II, III	Any	N1	Any	IVAD _o \times 4 cycles $+$ IVA \times 5 cycles $+$ maintenance \times 6 cycles	Yes	Yes

Note: Pathology (histology): *Favorable* indicates all embryonal cells, spindle cells, botryoid rhabdomyosarcoma (RMS); *unfavorable*, all alveolar RMS (including the solid‐alveolar variant). Postsurgical stage (according to the Intergroup Rhabdomyosarcoma Studies [IRS] grouping): *Group I* indicates primary complete resection (R0); *group II*, microscopic residual (R1) or primary complete resection but N1; *group III*, macroscopic residual (R2). Site: *Favorable* indicates orbit, genitourinary nonbladder prostate (i.e., paratesticular and vagina/uterus), and nonparameningeal head and neck; *unfavorable,* all other sites (parameningeal, extremities, genitourinary bladder‐prostate, and *other* sites). Node stage (according to the tumor‐node‐metastasis [TNM] classification): *N0*, no clinical or pathologic node involvement; *N1*, clinical or pathologic nodal involvement. Size and age: *Favorable*, tumor size (maximum dimension) ≤5cm *and* age younger than 10 years; *unfavorable*, all others (i.e., tumor size >5 cm *or* age 10 years and older).

Abbreviations: IVA, ifosfamide, vincristine, and dactinomycin; IVADo, ifosfamide, vincristine, and dactinomycin plus doxorubicin; VA, vincristine and dactinomycin.

Table S1). Ninety‐four of 362 patients (26%) who had ARMS were fusion-negative, a result that is consistent with the literature.

The protocol encouraged histology to be centrally reviewed, and 73% of patients had their diagnosis reviewed by a national reference pathologist and/or by the international E*p*SSG Pathology Panel. *FOXO1* fusion status assessment was performed in different laboratories according to national arrangements. It was not mandatory, and treatment usually was given based on the histopathology diagnosis. The analysis presented here was performed according to the final diagnosis; i.e., the diagnosis reviewed centrally or, if this was missing, the local diagnosis.

Patients were assigned to one of the four RMS 2005 risk groups and were treated according to the protocol guidelines that were previously described in detail $7-10$ and are summarized in Table 1. The protocol included two randomized trials for high‐risk patients that evaluated: (1) two regimens of chemotherapy in the first part of treatment: ifosfamide, vincristine, and dactinomycin plus doxorubicin (IVA) versus the IVA plus doxorubicin (IVADo) regimen (IVA plus doxorubicin 30 mg/m² on days 1 and 2 in the initial four cycles of chemotherapy followed by five cycles of IVA); and (2) the addition of a maintenance treatment with low‐dose cyclophosphamide and

vinorelbine for patients who were in clinical complete remission after initial standard treatment.

Delayed surgery and/or radiotherapy were planned after assessing tumor response to the three initial cycles of chemotherapy. When a residual mass was identified, surgical resection was encouraged if clear margins were achievable without organ or functional impairment. Marginal resection was acceptable at sites where complete resection was deemed unfeasible provided it was followed by radiotherapy. Radiotherapy was the only local treatment for patients who were not suitable for secondary surgery because of the tumor's location (i.e., parameningeal RMS). Radiotherapy doses were delivered according to histology, chemotherapy response, and surgical results: 41.4 grays (Gy) were given to patients who had ARMS in IRS group I or II, to patients in IRS group III who achieved complete remission after secondary surgery, and to patients with embryonal RMS who achieved complete remission with initial chemotherapy; and 50.4 Gy for patients with incomplete or unfeasible secondary resection. A boost of 5.4 Gy to the residual tumor was recommended for large tumors that responded poorly to chemotherapy. Radiotherapy to the involved lymph node sites was recommended at a dose of 41.4 Gy independent of histology and surgical resection.

Treatment was delivered with megavoltage photons, one fraction per day, 5 days per week, with conventional fraction sizes of 1.8 Gy per day.

This study was approved by the ethics committees of the participating centers, and informed consent was obtained from all patients according to the Declaration of Helsinki.

Statistical analysis

The primary end point was EFS, which was assessed by the investigator at each center and defined as the time from the date of study enrollment to the date of the first event, including death from any cause, disease progression (for patients who never achieved complete tumor remission), relapse after previous complete remission, the appearance of a new tumor, or the time of the latest follow‐up.

The following baseline variables were considered for the multivariable model: age at diagnosis, sex, histology, primary tumor site, IRS group, tumor size, nodal status, and *FOXO1* fusion status. For the purpose of this analysis, histology was maintained with its original clinical classification as either *favorable* (embryonal, spindle cell) or *unfavorable* (ARMS and not otherwise specified), whereas categorization for age, tumor size, and primary tumor site was re‐evaluated to confirm or to establish new groups with favorable and unfavorable prognoses. The treatment received was not included as an independent prognostic factor because it was administered according to clinical patient characteristics that were considered in the multivariable model.

No formal sample size was calculated because we used an event‐ per-candidate variable for the derivation of the model.^{[11](#page-9-0)} Patient characteristics were summarized as the median and interquartile range for continuous variables, or as the count and percentage for categorical variables. To evaluate the functional form of age and tumor size, these continuous variables were plotted against martingale residuals of a null Cox proportional hazards model. Cutoff values were determined based both on a visual evaluation of martingale residual distribution and on cutoff points corresponding to the most significant relation with the risk of event, estimated with a maximally selected log-rank statistic for values between the 10% and 90% quantiles using the upper bound of the *p* value, as described by Hothorn and Lausen, 12 as well as on optimal equal-hazard ratio method to discretize a continuous variable that has a U‐shaped relationship with log relative hazards in survival data. 13 The classification for primary tumor site was defined both on a visual evaluation of the martingale residual distribution and on a pairwise log‐rank test with Benjamini–Hochberg correction.

Median follow‐up was computed using the reverse Kaplan–Meier method. The occurrence of second‐order interactions was verified using a likelihood ratio test comparing models with and without the interaction terms.

Main effects and significant second‐order interactions of candidate predictors were included in a multiple Cox proportional hazards regression model. No deviation from the proportional hazards

assumption was found by the test statistic of Grambsch and Therneau.[14](#page-9-0) A backward elimination with the Akaike information criterion was applied for selecting all independent prognostic variables. A nomogram of the final reduced Cox regression model was generated for predicting 3‐year and 5‐year EFS probabilities. Model performance was evaluated by examining measures of discrimination and calibration. Discrimination (i.e., the ability of the model to differentiate between high‐risk and low‐risk patients) was calculated using the Harrell concordance (C) index, adjusted through 1000 bootstrap resamples. Bias‐corrected calibration plots at 3‐year and 5‐year EFS rates were produced by using a bootstrap procedure (1000 resamples) to account for consistency between observed and estimated survival probabilities.

Patients were stratified into four risk groups based on their individual score in the nomogram corresponding to optimal cutoff points, 12 and the log-rank test was used to compare groups. The survival probabilities were estimated using the Kaplan–Meier method and were reported with their 95% confidence interval (CI), which was calculated according to log-log transformation.

Statistical analyses were performed using R version 4.0.2 (R Foundation for Statistical Analysis) and the R packages rms, survival, survminer, and ggplot2.

RESULTS

Patient characteristics and outcome

Patient characteristics are summarized in Table [2.](#page-4-0) The median follow‐up was 6.3 years (interquartile range, 4.5–8.6 years). During follow‐up, 491 patients had an event, and 326 died. Overall, 245 (50%) locoregional relapses, 128 (26%) metastatic progressions, 84 (17%) progressive diseases, 24 (5%) second malignancies, five (1%) deaths from disease, and five (1%) failures because of toxicity were registered. The pattern of treatment failures appeared different across age groups, with higher local failures in younger children: 24% in patients younger than 3 years, 18% in the group aged 3–10 years, and 19% in the older patients ($p = .0156$; see Table S2).

The EFS and overall survival rates at 5 years were 71% (95% CI, 69–73 years) and 81% (95% CI, 79–83 years), respectively.

Continuous variable categorization

Considering age as a continuous variable, the risk of an event was greater for patients younger than 3 years, it decreased moving toward age 10 years, plateaued from ages 10 to 14 years, and was higher again in older patients (see Figure S1). Therefore, we categorized age at diagnosis as follows: younger than 3 years, 3–9 years, and 10 years or older.

When tumor size was considered as a continuous variable, the risk of an event increased when the greatest tumor dimension was >5 cm (see Figure S1). Similarly, the single estimated cutoff point that

TABLE 2 Patient characteristics.

Abbreviations: CI, confidence interval; HR, hazard ratio; IRS, Intergroup Rhabdomyosarcoma Studies; Q, quartile.

^aData were available for 1445 patients.

corresponded to the most significant relationship with outcome was 5 cm.

The analysis of the primary tumor site identified three groups with different risks of an event: (1) bile duct and genitourinary nonbladder/prostate sites; (2) orbit, head and neck nonparameningeal, and bladder/prostrate sites; and 3) extremity, parameningeal, and other sites (see Figure S1).

Multivariable model

Interactions involving age at diagnosis with tumor primary site, tumor size and *FOXO1* fusion status, sex with tumor primary site, and

IRS group with tumor primary site (see Table S3) were identified as significant and were included in the multivariable Cox regression model. A backward elimination procedure based on Akaike information criterion (6924.83) in the multivariable modeling retained five prognostic factors, including age at diagnosis interacting with tumor size, tumor primary site, IRS group, and *FOXO1* fusion status (Table [3\)](#page-5-0). The nomogram predicting 3‐year and 5‐year EFS is pre-sented in Figure [1.](#page-5-0)

The calibration plot for internal validation (see Figure S2) showed good agreement of 3‐year and 5‐year EFS probabilities between the estimated outcomes and actual observations. The C‐index was 0.66 in the original data, and the optimistic-corrected C statistic with 1000 bootstrap replications was 0.65.

Based on each patient's total score in the nomogram, patients were stratified into four groups: a low-risk group (176 of 1661 patients [11%]; total score, <68), an intermediate‐risk group (701 of 1661 patients [42%]; total score \geq 68 and total score <182.4), a highrisk group (423 of 1661 patients [26%]; total score \geq 182.4 and total score <232), and a very-high-risk group (361 of 1661 patients [22%]; total score, ≥232). The 5‐year EFS rates were 94%, 78%, 65%, and 52%, respectively, in the low-risk, intermediate-risk, high-risk, and very-high-risk groups (Table [4](#page-6-0) and Figure [2A](#page-7-0)); and the corresponding 5‐year overall survival rates were 97%, 92%, 74%, and 61% (Figure [2B](#page-7-0)).

DISCUSSION

There is a continuous need to refine risk classification for pediatric tumors to confirm the prognostic variables used in the past and incorporate new findings as they are discovered and to help choose the best possible treatment for each patient. This analysis represents an effort to review the E*p*SSG classification, which has been in use since 2005 and has served as a basis for the current E*p*SSG FAR‐RMS trial. It is particularly important also to try to incorporate molecular findings into a classification system that, to date, has been based essentially on clinical factors.

A recent study validated the clinicopathologic factors used in the COG studies and confirmed that patients older than 10 years, unfavorable tumor site, and unfavorable tumor size (>5 cm) are associated with an inferior outcome. Clinical group, nodal involvement, and histology were also confirmed as prognostic factors.^{[3](#page-9-0)}

The E*p*SSG adopted the same factors in the RMS 2005 study, but they were combined in a different way and determined a different treatment allocation for at least a proportion of patients with RMS. Our analysis confirms the prognostic value of most of the factors we used previously, but it also presents important new information.

The role of patient age as a prognostic variable is difficult to establish because the biologic characteristics and treatment modal-ities applied may change depending on age.^{[15](#page-9-0)} The unfavorable fusionpositive ARMS is more common in older children, and the favorable spindle cell VGLL2/NCOA2‐positive RMS is typical of infants. Conversely, the treatment of younger children is challenging, and the RMS 2005 protocol recommended age–dose adaptation of **TABLE 3** Multiple Cox regression model for event-free survival and nomogram coefficients.

Abbreviations: CI, confidence interval; E, events; GUBP, genitourinary bladder‐prostate; HNnoPM, head and neck nonparameningeal; HNPM, head and neck parameningeal; HR, hazard ratio; IRS, Intergroup Rhabdomyosarcoma Studies; Ref, reference category.

FIGURE 1 A nomogram for predicting 3‐year and 5‐year EFS probabilities. EFS indicates event‐free survival; GUBP, genitourinary bladder‐prostate; HNnoPM, head and neck nonparameningeal; IRS, Intergroup Rhabdomyosarcoma Studies Group.

chemotherapy for infants and a case‐by‐case discussion to decide on the use of radiotherapy in children younger than 3 years. Therefore, it is not surprising to find that children younger than 3 years have a relatively poorer prognosis. The relatively higher proportion of local failures in this group may be determined by difficulties with implementing an aggressive local treatment, and particularly radiotherapy, because of concerns about late sequelae.

The upper cutoff of 10 years and older is currently used both in E*p*SSG and COG trials, and it is confirmed by our analysis. Because the risk of failure is similar in patients aged 10–14 years, an age limit of 14 years could also be considered to identify patients at higher risk.

We confirmed that a tumor size of 5 cm in greatest dimension is the optimal cutoff to separate children with differing risk. This variable is easy to use, and it is not clear whether considering two or

TABLE 4 Patients' characteristics according to risk group.

Abbreviations: CI, confidence interval; EFS, event‐free survival; GUBP, genitourinary bladder‐prostate; HNnoPM, head and neck nonparameningeal; HNPM, head and neck parameningeal; IRS, Intergroup Rhabdomyosarcoma Studies; OS, overall survival.

three dimensions of the tumor may be more appropriate (but it is certainly more complicated).^{[16](#page-9-0)}

However, in our analysis, and as shown previously, age and tumor size outcomes were interdependent, confirming that older children who have large tumors represent the population at higher risk of treatment failure.

Compared with patients who have RMS arising in the extremities, parameningeal sites, or in the so‐called *other* sites, those located in the bladder/prostate and bile ducts had better outcomes. The latter sites were included in the unfavorable group in the RMS 2005 study; however, in light of the good results obtained in RMS 2005, we decided to

move them into the more favorable standard‐risk group in the FaR‐ RMS trial. For RMS arising in a biliary site, this is in contrast with the results presented by theCOG,which recently decided to include biliary RMS in the unfavorable category because of the suboptimal outcomes of patients treated in low-risk studies.^{[17](#page-9-0)} This difference may be explained in part by the different dose of alkylating agents administered to this group of patients in the E*p*SSG and COG studies and demonstrates the necessity of a common analysis and classification.

Clinical group has been identified as a major prognostic determinant since the initial cooperative studies on $RMS²$ and it has always retained its value.

FIGURE 2 Kaplan–Meier curves for (A) event-free survival and (B) overall survival stratified by risk group.

FOXO1 fusion status has been identified as an independent prognostic factor in several retrospective studies. In a recent analysis published by the COG, only the presence of metastases surpassed *FOXO1* fusion status as a prognostic factor.^{[6](#page-9-0)} This led the COG to include *FOXO1* status in their stratification system. We present similar results in this report, further supporting the use of *FOXO1* status rather than histology to assign treatment. The inclusion of *PAX3/7‐FOXO1* fusion status in risk stratification in place of histology in the FaR‐RMS protocol represents a first attempt to include tumor molecular characteristics in a risk‐ stratification system. Additional prognostic biologic factors have been identified in RMS. *MYOD1* and *TP53* mutations have been associated with a worse prognosis, whereas NCOA2/*VGLL2*‐associated gene fusions have a very good prognosis. The RMS 2005 study was not designed to collect these data in our population, so their inclusion in the E*p*SSG stratification system is under debate. These and other new biologic factors could have a very important role in stratifying patients.

The independent role of nodal involvement has been controversial. A lower survival rate has been reported both in patients with ARMS and in those with node‐positive embryonal RMS included in the RMS 2005 study. $8,18$ It is possible, however, that other tumor characteristics may be more important when a more intensive treatment is adopted. In addition, the gradual introduction of more sensitive imaging methods like fluorodeoxyglucose‐ positron emission tomography may have changed the evaluation of nodal involvement, possibly upstaging patients who have a lower tumor load and a better prognosis.

The impact of treatment has not been included in our model, and this represents a limitation of our study. However, treatment is determined by the risk group assigned to the patient on the basis of the initial disease and patient characteristics. Therefore, we mainly aimed to identify prognostic factors that can stratify patients at diagnosis.

In addition to identifying the role of different prognostic factors, the merit of this analysis is the production of a nomogram that may be used to calculate the prognosis for each patient based on currently known risk factors. This methodology is in use for adult patients with sarcoma. Validated nomograms can be used to predict overall survival and distant metastases in patients after surgical resection of soft tissue sarcoma of the extremities.¹⁹ A nomogram has been developed to estimate the chance of salvage for individual children with relapsed RMS treated according to the International Society of Pediatric Oncology Malignant Mesenchymal Tumor protocols to direct therapy appropriately toward cure, use of experimental therapies, and/or palliation.[20](#page-9-0)

The nomogram we propose is based on a study analyzing d a large, international, multicenter population that was treated homogeneously and had data prospectively collected. It can be used at diagnosis to assist clinicians in guiding treatment. At the same time, it should be recognized that we have not yet undertaken an external validation, which could be facilitated through an international collaboration. Moreover, it should be noted that this analysis used data from a prospective study in which the assessment of fusion status was not mandatory.

In conclusion, the results presented provide the rationale to modify the E*p*SSG stratification for adoption in the current FaR‐RMS trial, confirming the adverse prognostic value of an age of 10 years and a tumor size of 5 cm. It also supports reconsideration of the role of primary tumor site. Bladder/prostate and biliary tree RMS are now included in the favorable category in the FaR‐RMS trial. The most significant change is probably represented by the replacement of histology with fusion status. This makes the E*p*SSG stratification more similar to that of the COG system and will facilitate data comparison in the future.

In the meantime, the international community has recognized the need to adopt a common stratification system. This represents the main goal of the recently established INSTRuCT consortium^{[21](#page-9-0)} and will help establishing a common language and, hopefully, risk stratification in RMS treatment and research.

AUTHOR CONTRIBUTIONS

Gian Luca De Salvo: Conceptualization, writing–original draft, writing–review and editing, and methodology. **Paola Del Bianco**: Conceptualization, methodology, formal analysis, writing–original draft, and writing–review and editing. **Veronique Minard‐Colin**: Conceptualization, investigation, funding acquisition, writing–review and editing, and data curation. **Julia Chisholm**: Conceptualization, data curation, investigation, funding acquisition, and writing–review and editing. **Meriel Jenney**: Conceptualization, data curation, investigation, writing–review and editing, and funding acquisition. **Gabriela Guillen**: Data curation and investigation. **Christine Devalck**: Data curation and investigation. **Rick Van Rijn**: Investigation. **Janet Shipley**: Investigation. **Daniel Orbach**: Data curation and investigation. **Anna Kelsey**: Investigation. **Timothy Rogers**: Investigation. **Florent Guerin**: Investigation. **Giovanni Scarzello**: Investigation. **Andrea Ferrari**: Investigation and data curation. **Maja Cesen Mazic**: Data curation and investigation. **Johannes H. M. Merks**: Conceptualization, data curation, investigation, and writing–review and editing. **Gianni Bisogno**: Conceptualization, data curation, investigation, writing– original draft, writing–review and editing, and funding acquisition.

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CONFLICT OF INTEREST STATEMENT

Veronique Minard‐Colin reports grants from F. Hoffmann‐La Roche outside the submitted work. Julia Chisholm reports service on a Children's Oncology Group Data Safety and Monitoring Board and support for other professional activities from the National Cancer Institute outside the submitted work. Meriel Jenney reports personal/consulting fees from the Cardiff and Vale University Health Board outside the submitted work. Daniel Orbach reports personal/ consulting fees from Bayer Healthcare, EUSA Pharma (US) LLC, and F. Hoffmann‐La Roche; and service on a Data Safety and Monitoring Board for Eli Lilly and Company outside the submitted work. Johannes H. M. Merks reports personal/consulting fees from Bayer, GlaxoSmithKline, and Merck outside the submitted work. Gianni Bisogno reports personal/consulting fees from GlaxoSmithKline and support for other professional activities from Bayer outside the submitted work. The remaining authors disclosed no conflicts of interest.

DATA AVAILABILITY STATEMENT

The data that support the findings of this study are available from the corresponding author upon reasonable request.

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SUPPORTING INFORMATION

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