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Divergent age-related changes in parasite infection occur independently of behaviour and demography in a wild ungulate

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Abstract

As animals age they exhibit a suite of phenotypic changes, often including reductions in movement and social behaviour ("behavioural ageing"). By altering an individual's exposure to parasites, behavioural ageing may influence infection status trajectories over the lifespan. However, these processes could be confounded by age-related changes in other phenotypic traits, or by selective disappearance of certain individuals due to parasite-induced mortality. Here, we uncover contrasting age-related patterns of infection across three helminth parasites in wild adult female red deer (*Cervus elaphus*). Counts of strongyle nematodes (order: Strongylida) increased with age, while counts of liver fluke (*Fasciola hepatica*) and tissue worm (*Elaphostrongylus cervi*) decreased, and lungworm (*Dictyocaulus*) counts did not change. These relationships could not be explained by socio-spatial behaviours, spatial structuring, or selective disappearance, suggesting behavioural ageing is unlikely to be responsible for driving age trends. Instead, social connectedness and strongyle infection were positively correlated, such that direct age-infection trends were directly contrasted with the effects implied by previously documented behavioural ageing. This suggests that behavioural ageing may reduce parasite exposure, potentially countering other age-related changes. These findings demonstrate that different parasites can show contrasting age trajectories depending on diverse intrinsic and extrinsic factors, and that behaviour's role in these processes is likely to be complex and multidirectional.

30 Introduction

31 An individual's disease status depends on a combination of its exposure and susceptibility to parasites
32 [1,2]. Exposure is broadly a function of an individual's social and spatial behaviour within the context
33 of a population, and between- and within-individual variation in behaviour can have important
34 consequences for infectious disease status [3–5]. In humans and wild animals, individuals alter their
35 behaviour as they age [6,7], with a series of general changes characterised by reduced movement [8–
36 10] and sociality [10,11], and specifically a tendency towards positive interactions with specific
37 individuals known as "social selectivity" [11–14]. Because these processes influence the way that
38 individuals contact each other, they could influence rates of exposure to pathogens and therefore their
39 infection probability [6]. Nevertheless, because behaviour has yet to be linked to age-related changes
40 in infection in a given population, the role of behavioural ageing in driving infection dynamics remains
41 unclear [6].

42 Behaviour could drive age-related changes in infection status through a series of mechanisms (see [6]
43 for a review). For example, individuals could alter their feeding locations as they age, which could move
44 them into areas that are more or less likely to support environmental parasites [5], or it could result in
45 lower-quality resource intake, driving weaker immunity and therefore greater susceptibility to infection
46 [15]. Similarly, ageing individuals could become more socially isolated, potentially driving decreased
47 exposure to directly transmitted parasites [7,10]. More subtly, if ageing individuals tend to prefer a few
48 close associates over socialising broadly (i.e. showing increasing selectivity), this could drive an increase
49 in modularity, with complex outcomes for epidemiological dynamics [6,16]. However, age also alters
50 other phenotypic and demographic changes that could complicate these relationships. For example,
51 ageing individuals experience a suite of physiological changes [17], many of which affect the immune
52 system (i.e., "immunosenescence" [18–20]). Because these changes often result in increased
53 susceptibility to infection, the conventional wisdom is that individuals will exhibit a greater prevalence
54 or burden of parasites as they senesce [20–22]. Alternatively, individuals may acquire adaptive
55 immunity to certain parasites as they become exposed, potentially leading to an increase in immunity
56 to these particular parasites [23]. Additionally, because parasites often exact survival costs on their
57 hosts, more heavily-infected individuals may be more likely to die – a process known as "selective
58 disappearance" – which could produce a negative age-infection trend at the population level, and may
59 bias estimates of within-individual ageing patterns [24,25]. The emergent pattern of infection status
60 over the lifespan will depend on a combination of these factors.

61 Given these combined behavioural, immunological, and demographic changes, ageing individuals'
62 infection statuses could be asynchronous and divergent for different parasite taxa, leading to an age-
63 related shift in parasite community composition. This possibility is supported by the literature on
64 observed age-infection relationships, which comprises a wide diversity of positive, negative, and
65 nonlinear changes in prevalence and intensity of infection (e.g. [26–31]); however, most such studies

66 focus on one parasite taxon, and it is therefore unclear how often parasites show divergent age-related
67 trends within a population. These studies are likewise often cross-sectional rather than longitudinal (i.e.
68 they do not follow the same known individuals through time), and are therefore unable to identify and
69 extricate selective disappearance effects ([24,25]; but see [26,27,30]). This is an especially important
70 gap in our understanding, particularly given that parasites are generally defined by their ability to cause
71 harm to their hosts [6,32] and will therefore likely drive patterns of disappearance. Additionally, because
72 studies rarely model variable age-infection relationships within a given population (which requires
73 longitudinal data), it is unclear how these drivers drive variable infection trajectories over the lifespan
74 [6]. Contrasting age trajectories for different parasites in the same individuals within the same
75 population may help to untangle the mechanisms underlying age-infection trends more broadly.

76 Here, we ask how different helminth parasite counts change over the lifespan in a long-term study
77 population of wild red deer (*Cervus elaphus*), in which female deer are monitored from birth until their
78 death, generally at least a decade later. Building on a rigorous behavioural censusing operation [33] in
79 a society with well-understood spatial structuring [34], studies have shown that female deer strongly
80 alter their foraging and social behaviour as they age [9,10]. Specifically, they reduce their home range
81 sizes [9] as well as moving towards areas of lower density at the periphery of the population, and
82 become less socially-connected [10]. The deer also feature high-resolution individually-tied egg and
83 larval counts of multiple helminth parasite taxa. These parasites infect individuals throughout their lives
84 without inducing full immunity, and therefore exist at high prevalence in the population, but with
85 substantial within- and between-individual variation that enables testing of a wide range of ecological
86 questions [35]. Counts of these parasites fluctuate seasonally [35] and are influenced by allocation of
87 resources to reproduction [36], as well as having strong costs in terms of survival and reproduction
88 [37]. Combining these sources of information, this population is well-suited to examining long-term age
89 trajectories of infection by multiple parasites, and the possible role of behaviour and demography in
90 driving them. Specifically, we ask 1) how counts of multiple helminth parasites change over the lifespan;
91 2) whether these counts are influenced by spatial and social behaviours governing rates of exposure;
92 and 3) whether these behaviours could explain or counteract the age-related changes we see.

93 Methods

94 Study population

95 The study population was the individually-monitored Isle of Rum red deer; this unmanaged wild
96 population has been studied since 1973 [33], with regular faecal parasite sampling since 2016 [35].
97 The deer are censused 40 times a year, with individuals known by name and individually marked using
98 a combination of coloured and patterned collars, tags, and ear punches. When identified in a census,
99 an individual's location (to the nearest hectare) is recorded, providing it with an easting and northing
100 location in two-dimensional space; further, groups of deer are identified in the course of censusing and

101 taken by the field worker to be associating, forming the basis for the social network pipeline described
102 below. The vegetation type each deer is on is noted.

103 The deer give birth in May and June, and daily censuses over the calving period allow >90% of calves
104 to be caught, tagged, and weighed. The deer year runs from 1st of May, and individuals are assigned
105 an age in years based on the deer year they were born in; for example, all individuals turn 1 year old
106 on the 1st of May the year after they were born. 40 study area censuses per year allow us to keep track
107 of each individual's life history, and individuals have known death dates, generally to within one month,
108 and often to the day, allowing accurate quantification of mortality. Following our previous related work
109 in this system [10,34], here we assess mature females (3 years and older), as these are the best-
110 understood age and sex class, with the largest available dataset; young males disperse and few adult
111 males live in the study area, and so males are less well sampled. Female reproductive status in any
112 year was coded as either "none" (did not give birth that year), "summer" (gave birth, but the calf died
113 before October 1st), or "winter" (gave birth, but the calf died over its first winter or reared it through
114 its first winter). This categorisation is based on the relative costs of reproduction, which are observed
115 to be high in individuals whose calf survives to the winter, regardless of whether the calf then survives
116 to the spring; these costs are reflected both in terms of parasitism and fitness [36–38].

117 Parasitology

118 We have previously described our parasitology monitoring regime in detail [35]. Briefly, three times a
119 year (late April, August, and November), for two weeks at a time, we observe the deer intensively to
120 collect faecal samples from as many individuals as possible. After observing an individual defaecating,
121 we collect the sample as soon as possible into a resealable plastic bag, and at the end of the day we
122 homogenise it, and store it anaerobically (i.e., with the bag sealed) in the fridge at ~4°C until counting.
123 By observing the individual and noting the location of the defaecation event itself, coupled with
124 collection within a short period (generally within an hour and most often within 10-20 minutes), we are
125 able to tie samples to known individuals.

126 We counted gastrointestinal helminth parasite propagules in these samples using a variety of
127 techniques. We counted strongyle nematode (order: Strongylida) eggs within three weeks of collection
128 using a salt flotation-centrifugation technique, where a gram of homogenised faecal matter was mixed
129 with saturated salt and the mixture homogenised, causing the eggs of a selection of parasites to rise
130 to the surface where they can be easily counted [35]. Liver fluke (*Fasciola hepatica*) eggs were counted
131 using a sedimentation technique, where a weighed amount of faecal matter was mixed with a large
132 amount of sediment and allowed to settle over three minutes, and the supernatant removed via vacuum
133 suction. Finally, tissue worm (*Elaphostrongylus cervi*) and lungworm (*Dictyocaulus* sp.) larvae were
134 counted using a Baermannisation technique, in which a weighed amount of faecal matter was wrapped
135 in porous cloth and submerged in water over 24 hours to allow the mobile larvae to escape, which were
136 then reduced in volume by vacuum suctioning and preserved for counting. All techniques are accurate

137 to at least 1 egg or larva per gram. These different assays were required because of the different
138 physical properties of the propagules: strongyle eggs float in saturated salt solution, whereas fluke
139 eggs are too heavy and must be sedimented, while tissue worm and lungworms are alive and possible
140 to isolate using their movement behaviour. Our salt flotation also detected a number of other parasites
141 (described in [35]), but they were present at low prevalence (<10%) in adult females, and therefore
142 we were less able to analyse how they changed with age.

143 Samples were collected between August 2016 and April 2021. Where multiple samples were collected
144 for a given individual in a given sampling trip, we took the mean of the counts to leave a maximum of
145 one count per individual per sampling trip. We did so because there were relatively few within-season
146 repeats, and they were restricted to the beginning of the study (N=654 repeats). Our final dataset
147 included Ns=1449 measurements taken from Ni=210 individuals; some assays were not completed for
148 all samples, leaving Ns=1433 *F. hepatica* measurements, and Ns=1126 *E. cervi* and *Dictyocaulus*
149 measurements taken from Ni=209 individuals. The number of samples per individual and per sampling
150 trip are displayed in Supplementary Table 2.

151 Behavioural metrics

152 We examined how an individual's behaviour was associated with its parasite burden. Building from our
153 prior findings that individuals alter a suite of socio-spatial behaviours as they age, we selected a series
154 of behaviours to test. All such behaviours are expected to influence some element of exposure to
155 parasites, involving either movement to different areas on the landscape or interactions with other
156 individuals. We used all census observations of each individual in each year, including adults and
157 juveniles. We chose to include juveniles in the social network as they are heavily infected with parasites
158 [35] and could therefore play an important role in infecting older individuals. The behavioural metrics
159 include:

160 **Social network metrics:** We constructed social networks as previously described [10,34]. Social
161 connections were judged by field workers based on a spatially-parameterised "gambit of the group"
162 approach, where individuals within a certain distance were taken to be socialising by the field worker
163 (see [10,34] for details), as described above. First, we took the average group size for each individual
164 across the year. Next, for each year, we constructed networks based on these associations, which we
165 corrected for observation bias using the simple ratio index [39] such that each dyad's connection was
166 scaled between 0 (never seen together) to 1 (never seen apart). We then calculated two network
167 metrics: degree centrality (i.e., the number of individuals an individual was seen with over the course
168 of a year), and strength centrality (i.e, the summed weighted connections to all individuals over the
169 year).

170 **Local population density:** We calculated local density using a previously described pipeline for this
171 population [10,34], using all observations of each individual in each year, including both adults and
172 juvenile individuals. This approach uses a kernel density estimator, taking individuals' annual centroids

173 and fitting a two-dimensional smooth to the distribution of the data, producing a two-dimensional
174 spatial distribution of the population. Individuals are then assigned a local density value based on their
175 location on this kernel.

176 **Spatial behaviour metrics:** we included several metrics that quantitatively described an individual's
177 spatial behaviour in the study area, all of which have been shown to change with age [10]. These
178 included: population centroid distance (the distance from the overall mean location of the population,
179 which increases with age); graze type (the proportion of sightings in which an individual was seen on
180 high-quality grazing, which decreases with age); and home range area (built based on each individual's
181 density distribution, which decreases with age).

182 **Time lag:** We examined how annual behaviour metrics from deer year t influenced parasite infection
183 in deer year $t+1$. To put this in terms of calendar years, we examined how an individual's behaviour
184 from May 1st in year t to April 30th in year $t+1$ affected its parasite burden in August year $t+1$, November
185 year $t+1$, and April in year $t+2$.

186 Although a relatively coarse annual measure of behaviour, individual-level repeatability of annual social
187 network positions is high [34], as is repeatability of annual measures of spatial fidelity and home range
188 size [9,40], and previous work has shown these measures to be ecologically relevant for individuals
189 [10,34]. Using the previous deer year's social network also allowed us to accommodate the time lag of
190 the influence of social connections on parasite burden (e.g. including parasites' time to development
191 and maturation and egg production, which generally take months to stabilise) and allowed us to avoid
192 confounding produced by analysing an individual's social connectedness in a given deer year with its
193 concurrent and earlier parasite infection status, and possible reverse causality emerging from e.g.
194 avoidance responses [41]. That is, including behavioural measures taken in deer year t in models
195 examining parasite infection through deer year t would involve including behavioural observations from
196 post-parasite sampling; because behaviours often change in response to infection, and often with
197 protective consequences that decrease the risk of infection [42], this could drive complex and
198 counterintuitive relationships with parasitism that we were not intending to test. Finally, behaviour of
199 the deer in this population is highly seasonal [33], as is parasite infection [35]; using sub-annual
200 measures of infection that differed between seasons might risk strong confounding between behaviour
201 and infection. As such, we judge our annual measures to be a reliable and parsimonious indicator of
202 social and spatial behaviour with relevance to the risk of parasite transmission over the lifespan.

203 Models

204 Our dataset included 1449 measures of parasite counts in 210 individual deer, spread across 5 deer
205 years and 15 collecting seasons. To identify age-related changes in parasite burden and determine how
206 they might arise, we fitted a selection of generalised linear mixed models (GLMMs) using the Integrated
207 Nested Laplace Approximation (INLA) in R [43]. INLA is a deterministic Bayesian algorithm that allows
208 fitting of spatially distributed random effects (Stochastic Partial Differentiation Equation effects, see

209 below) to account for spatial autocorrelation in the response variable [44]. All models were fitted with
210 uninformative default priors. Models were checked by simulating from the model posteriors and
211 inspecting the predicted against the observed values and examining them for uneven patterns. We
212 calculated P values from the posteriors using the `inla.pmarginal` function, providing the probability of
213 generating a result that overlapped with zero from the distribution. For all models, continuous predictors
214 were scaled to have a mean of 0 and a standard deviation of 1 before analysis. The model sets we
215 used were as follows:

216 **Base models:** first, we fitted models to understand individual age trajectories of parasitism in the
217 population. We examined each parasite count as a response variable with a negative binomial
218 specification, given their strongly overdispersed distribution. We fitted explanatory variables including
219 Year (factor with 5 levels: deer years? 2016-2020); Season (factor with 3 levels: Summer, Autumn, and
220 Spring); Reproductive Status (Factor with 3 levels: None, Summer, and Winter); Age (continuous
221 covariate, range 3-24, mean 7.9). We ran these models both without and with a random effect of
222 individual identity, to examine how controlling for among-individual variation impacted our estimates
223 of age effects. Using individual identity in this way can help to distinguish within-individual ageing
224 processes versus between-individual demographic processes [24]; fitting an ID effect and seeing the
225 disappearance of an age effect would imply that age was only associated with infection at the between-
226 individual level.

227 **Social models:** second, to identify the effects of a given behaviour on infection – and the effects of
228 incorporating said effect on age-infection relationships – we ran a series of models, each of which
229 added a behavioural metric to the base model. We then investigated the mean estimate and 95%
230 credible interval of this behavioural metric effect, and examined the impact that its inclusion had on
231 the age effect estimate to ask whether behaviour could be responsible. We fitted behavioural metrics
232 in a piecewise fashion – rather than adding them all at the same time – because the age-related
233 changes in behaviour with age tend to manifest as correlated socio-spatial behaviour syndromes [10].
234 We excluded counts from the autumn, because their low values precluded fitting as explanatory
235 variables in our models. Adding all at the same time would risk substantial collinearity, and fitting them
236 one at a time allowed us to test our hypotheses effectively.

237 **Spatial models:** for each model, to identify whether our results were affected by spatial
238 autocorrelation, we added a spatially distributed Stochastic Partial Differentiation Equation (SPDE)
239 effect [44–46] in INLA. This effect uses each individual's average annual easting and northing to model
240 how spatial proximity drove individuals to have similar parasite counts, according to Matern covariance.
241 Fitting this effect had several purposes: by comparing the fit of the spatial model with the base model,
242 we could identify whether the parasite counts were significantly spatially autocorrelated; by comparing
243 the model estimates we could identify whether this spatial autocorrelation was affecting our
244 conclusions; and by plotting the effect in space we could identify spatial hot- and coldspots of infection

245 [46]. To assess model fit, we used deviance information criterion (DIC), with a cutoff of -2 Δ DIC to
246 distinguish between competitive models.

247 **Survival models:** often, ageing models incorporate fixed effects of longevity to examine selective
248 disappearance of certain individuals [24]. We were unable to do this with our dataset, as it spanned
249 five years running to the present; because many individuals were yet to die, we did not have known
250 longevity values for many of the data points, which reduced our models' power in this context. As such,
251 to provide an approximate answer to this question, we fitted binomial survival models following previous
252 methodology [37] to examine whether parasites were likely to be causing annual mortality in adult
253 females (i.e., the same dataset we were testing for age-infection associations), and therefore might be
254 producing observed age-infection relationships. With observations from each individual:deer-year
255 combination as the unit of investigation, we fitted overwinter survival (0/1) as a response variable, with
256 explanatory variables including deer? Year; Reproductive Status; Age; and a random effect of individual
257 identity, all as described above. We sequentially added each parasite count (log(X+1)-transformed) as
258 an explanatory variable, one at a time, to investigate whether they correlated with subsequent survival.
259 In our dataset, there was an 89.1% annual survival rate across the 6 years of sampling; of our 208
260 individuals in the survival models, 83 (40%) died. We note that this is a relatively crude way of assessing
261 selective disappearance effects that was necessitated by our dataset; depending on the effects shown
262 by the mortality assessments, we may or may not be able to infer an effect of selective disappearance
263 using such an analysis. However, this approach to detecting survival effects has high statistical power
264 and has been used previously to detect strong survival effects of parasitism [37], which is the central
265 hypothesised cause of selective disappearance in this context; this paper expands on this analysis by
266 including more data, confirming the patterns using an expanded dataset and differently parameterised
267 models, and by testing multiple pathogens.

268 Results

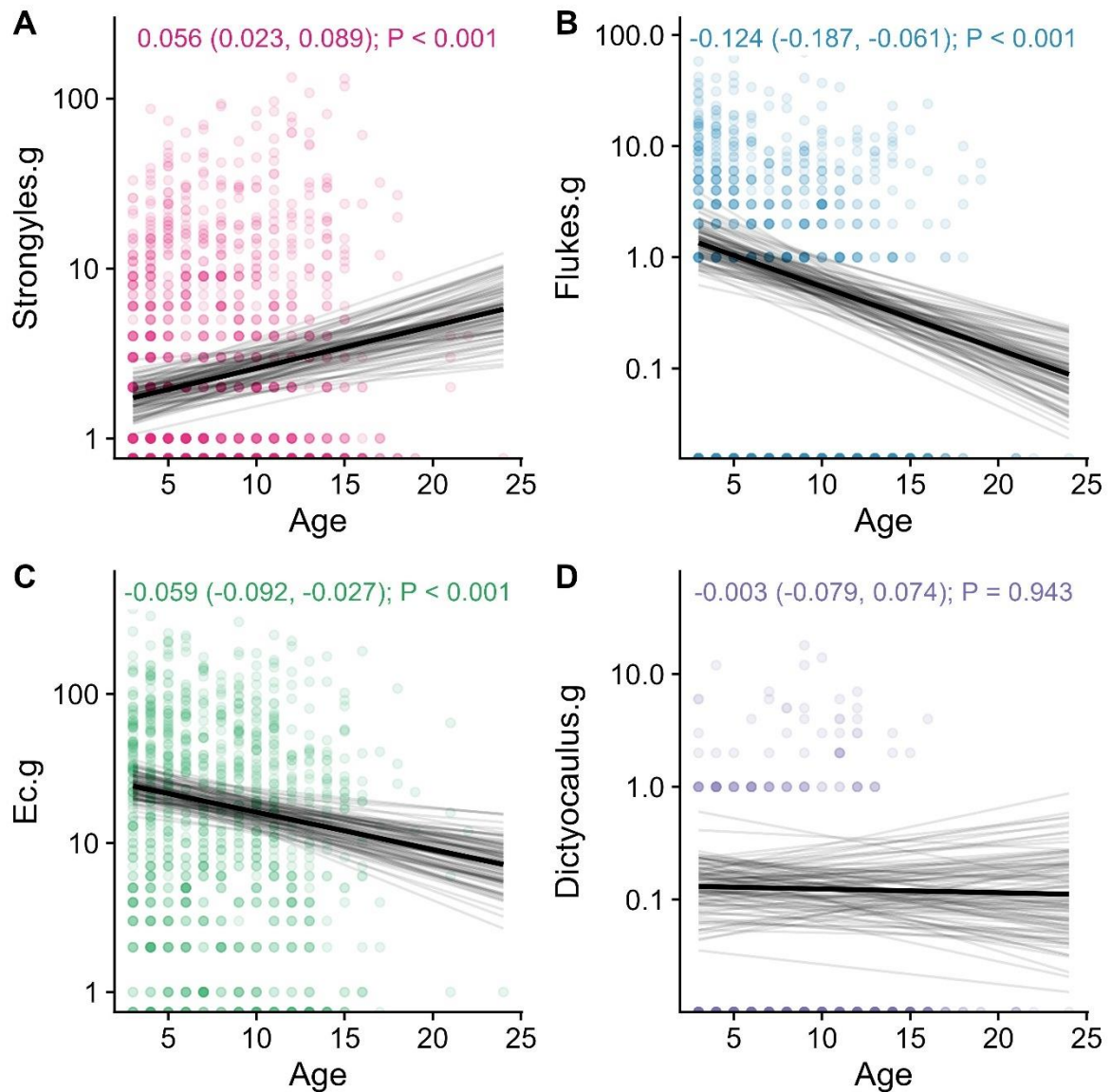
269 We found substantial contrasting age-infection relationships for three out of four parasites: there were
270 small positive associations between age and strongyle count (Figure 1A; mean effect estimate: 0.138,
271 lower 95% credibility estimate: 0.014, upper 95% credibility estimate: 0.261, $P=0.029$), and moderate
272 negative associations between age and liver fluke *Fasciola hepatica* (Figure 1B; 0.372, -0.605, -0.141,
273 $P=0.002$) and tissue worm (*Elphostrongylus cervi*) count (Figure 1C; -0.251, -0.364, -0.14, $P<0.001$).
274 *Dictyocaulus* lungworms, meanwhile, showed no relationship with age (Figure 1D; $P>0.05$). All effect
275 estimates and 95% credibility intervals are derived from the mean of the posterior effect distribution;
276 we report estimates here and in the model effects plots in units of standard deviations, but to aid
277 interpretation in the scale of the lifespan, in Figure 1-2 they are reported and displayed in units of age
278 in years or degree centrality respectively.

279 Spatial autocorrelation effects substantially improved the models for flukes and tissue worms
280 (Supplementary Table 1; Δ DIC<-3), but not for strongyles or lungworms (Supplementary Table 1;

281 $\Delta\text{DIC}>-2$). These findings demonstrate that there was notable heterogeneity in parasite infection
282 (Supplementary Figure 2), but controlling for this effect did not impact our age estimates (Figure 3A,
283 Supplementary Figure 1), demonstrating that changes in spatial location were unlikely to be responsible
284 for our observed age effects. There were moderate density effects evident in the base models for *E.*
285 *cervi* and *F. hepatica*, but these effects were removed when spatial autocorrelation was controlled for
286 (Supplementary Figure 1). The spatial distributions of these parasites largely agreed with earlier
287 observations [46], with greater *F. hepatica* count in the south-middle of the study area and greater *E.*
288 *cervi* count in a slow gradient moving towards the north, particularly the northeast (Supplementary
289 Figure 2).

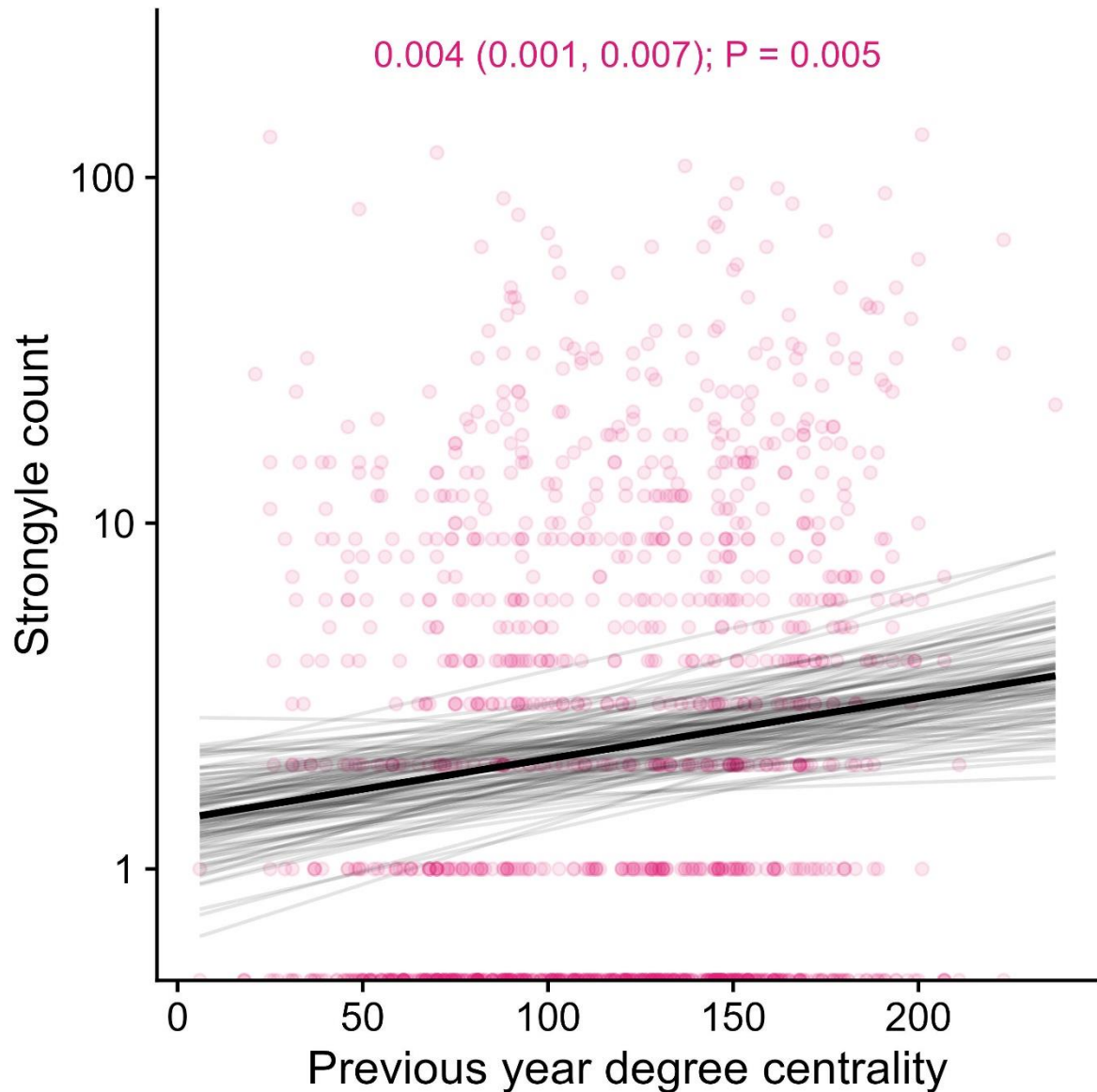
290 In our behavioural models, we uncovered a moderate positive effect of degree centrality on strongyle
291 infection (Figure 2; 0.171, 0.052, 0.289, $P=0.005$). There were a number of effects that were initially
292 significant in our non-spatial models but their effects were removed when spatial autocorrelation was
293 accounted for (Supplementary Figure 3), indicating that these behaviours were not possible to extricate
294 from spatial heterogeneity in the parasite's distribution. There was likewise a moderate effect of annual
295 density on *F. hepatica* infection, which persisted when spatial autocorrelation was controlled for (-
296 0.415, -0.796, -0.019, $P=0.04$; Supplementary Figure 3). In all cases, accounting for behaviours in the
297 models had very little impact on the age estimates (Supplementary Figure 3), demonstrating that age-
298 related changes in parasitism were largely independent of behavioural effects.

299 We found that strongyle count was strongly associated with reduced overwinter survival probability
300 (Figure 3B; Supplementary Figure 4; -0.98, -1.47, -0.55, $P<0.001$), agreeing with previous findings
301 [37]. This finding remained significant when spatial autocorrelation was controlled for (Supplementary
302 Figure 2). There were only weak negative nonsignificant trends with the other parasites (Figure 3B;
303 Supplementary Figure 2; $P>0.05$). Additionally, fitting random effects of individual identity substantially
304 improved model fit ($\Delta\text{DIC}<-10$; Supplementary Table 1) but without notably affecting the age effect
305 estimates (Figure 3A). Taken together, these findings provide little evidence for a role of selective
306 disappearance in *driving* our observations, except for potentially *obscuring* the age-strongyle trend.
307 That is, our estimate for the age effect on the strongyle counts is a composite that likely includes a
308 contrasting effect of selective disappearance, and is therefore likely an underestimate.



309

310 Figure 1: Age-related changes in infection with four helminth parasites in wild red deer. The x axis
 311 represents age in years. A) strongyle eggs per g; B) Fluke eggs per g; C), *E. cervi* larvae per g; D)
 312 *Dictyocaulus* larvae per g. Taken from the best-fitting models, the dark black line represents the
 313 mean of the posterior distribution for the age effect estimate; the light grey lines are 100 random
 314 draws from the posterior to represent uncertainty. The age effect estimate, credibility intervals, and P
 315 values are given at the top of each panel. The points represent individual samples, with transparency
 316 to allow for visualisation of overplotting. The y axis has been log₁₀-transformed; 0-counts (which are
 317 not possible to display on this logged scale) are displayed at the bottom of the graph.

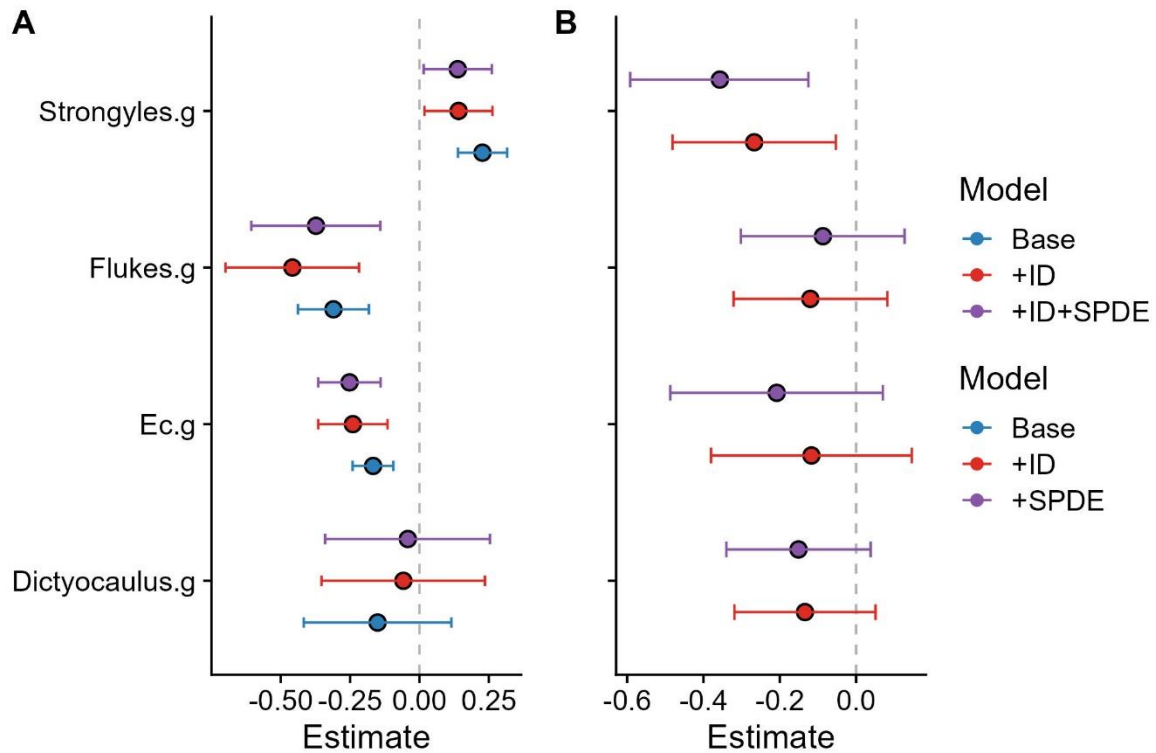


318

319 Figure 2: Association between social connectedness (degree centrality) in the previous year and
 320 strongyle nematode count in wild red deer. The x axis is in numbers of contacts; the y axis is in eggs
 321 per gram. Taken from the spatial model, the dark black line represents the mean of the posterior
 322 distribution for the age effect estimate; the light grey lines are 100 random draws from the posterior
 323 to represent uncertainty. The degree effect estimate, credibility intervals, and P values are given at
 324 the top of the figure. The points represent individual samples, with transparency to allow for
 325 visualisation of overplotting. The figure has been cropped to the distribution of the fitted lines to help
 326 visualising the model fits, so some points outside this range have been excluded from the figure.

327

328



329

330 Figure 3: Model effect estimates for A) the effect of age on parasite counts and B) the effect of
 331 parasite counts on overwinter survival probability. Points represent the mean for each effect estimate;
 332 error bars denote 95% credibility intervals. All estimates are given on the link scale, in units of
 333 standard deviations. Different colours represent different model constructions: "+ID" = includes a
 334 random effect of individual identity. "+SPDE" = includes the ID effect, plus an additional spatially
 335 distributed SPDE random effect to account for spatial autocorrelation.

336 Discussion

337 We uncovered substantial and contrasting age-related changes in parasite count across different
 338 parasites in a long-lived wild mammal, which were not explained through considering behavioural or
 339 demographic factors. Ageing red deer experienced a small increase in strongyle nematode counts,
 340 which contrasted with stronger age-related decreases in liver fluke (*F. hepatica*) and tissue worm (*E.*
 341 *cervi*) counts. These findings add to a sparse body of longitudinal individual-based evidence for age-
 342 related changes in parasite count in wild animals [20,21,26,27,30,47]. Accounting for and quantifying
 343 spatial autocorrelation and fitting socio-spatial behavioural metrics in our models had no detectable
 344 effects on our age estimates, suggesting that these changes were unlikely to be driven by previously
 345 documented behavioural ageing patterns and resulting changes in exposure rate [10]. Similarly, there
 346 was no evidence that selective disappearance of certain individuals was driving our observed trends,
 347 given that survival costs were limited to strongyle infection and were insufficient to produce our
 348 observed trends. As such, these observations do not suggest that behavioural ageing drives age-related

349 changes in parasite infection in this system, and instead imply that divergent age-related trends may
350 arise for different parasites through changes in intrinsic (e.g. physiological or immunological) traits.

351 Our observation that greater social connectedness predicted greater strongyle count agrees with the
352 conventional wisdom that infectious disease is a primary cost of sociality [48,49], but this trend was in
353 the opposite direction to the direction we expected if social behaviour was playing a role in driving age-
354 infection relationships. That is, if individuals' ageing behaviour were driving the effect, because social
355 connectedness decreases with age [10], we would expect strongyle count to likewise decrease with
356 age. Instead, these findings are more suggestive of the reverse: ageing individuals may reduce their
357 exposure to parasites as they decrease their social connectedness, which could ultimately minimise the
358 effects of a waning immune system for strongyles. Indeed, this mechanism has been theorised several
359 times [7,10], and recently received strong support via behavioural simulations [50]. If behavioural
360 ageing is linked to reducing exposure due to immunosenescence, because the strength of natural
361 selection is expected to wane in later life [51,52], it is unlikely that this is an adaptive response
362 specifically brought on by immunosenescence; instead, a relationship between behavioural ageing and
363 infection could emerge through more general behavioural compensation for a weak immune response
364 that evolved in earlier life and persists as the animal senesces. Such behavioural compensation is
365 relatively common [53,54]: for example, Stephenson [55] demonstrated that guppies (*Poecilia*
366 *reticulata*) show stronger conspecific avoidance when they are more susceptible to infection. Although
367 it has yet to be shown that immunosenescence and social ageing are linked directly, our observations
368 are consistent with a similar underlying process for strongyles. Conversely, although we noted a
369 negative correlation between density and *F. hepatica* infection, and individuals tend to move to areas
370 of lower density as they age [10], there was nevertheless a decrease in *F. hepatica* count with age.
371 Therefore behaviour was likewise countering age-related changes – but in the reverse pattern, by
372 potentially driving *greater* exposure to *F. hepatica* – which were nevertheless counteracted by other
373 phenotypic changes. Taken together, these findings indicate that behaviour likely plays a plastic or
374 buffering role mediating relationships between phenotypes, age, and infection as an individual ages.

375 It was also surprising that degree centrality – a social network metric – predicted strongyle count,
376 rather than any spatial behaviour metrics. This effect was relatively strong, and corresponded to roughly
377 a doubling in strongyle count across the range of degree centrality values (Figure 2). This was perhaps
378 unexpected as helminth parasites transmit indirectly, so we would expect that incorporating spatial
379 measures (rather than more direct measures of social contact) may be more representative of indirect
380 contact rates – and therefore of parasite counts [5]. For example, areas of higher density should be
381 more intensely used and therefore support greater larval concentrations on the pasture. Further, the
382 spatial autocorrelation effects in the models should account for age-related movements towards areas
383 of variable transmission of certain parasites – for example, if lower *F. hepatica* counts were driven by
384 movements away from wetter areas that tend to support transmission via their water snail intermediate
385 hosts [46]. Because social connections are parameterised according to spatiotemporal coincidence (i.e.,

386 they require individuals to be in the same location at the same time), the measures derived from this
387 metric could be more indicative of between-individual helminth transmission, which could occur more
388 on the timescale of days to months than years, even despite the fact that both social and spatial
389 behaviours were ultimately summarised at the annual level. Regardless of the ultimate cause, these
390 findings agree with the previous observation that social network position is both heavily intertwined
391 with spatial behaviour in this system and a biologically important stand-alone measure [10,34]. This
392 finding adds notably to the literature on spatial-social analysis in disease ecology, and accentuates the
393 value of using both spatial and social metrics when quantifying the drivers of infection status [5].

394 Aside from behaviour, a variety of age-related changes could be responsible for divergent age trends
395 among parasite taxa: on the immune side, increasing strongyle counts could be driven by decreased
396 resistance brought about by immunosenescence, agreeing with previous observations in wild Soay
397 sheep [21,56]. This observation disagrees with a previous finding that strongylid infection decreases
398 with age in African elephants, for example [31]; given that that investigation occurred at the population
399 level, it is possible that selective disappearance may have played a role in influencing this pattern in
400 the elephants, accentuating the benefit of longitudinal individual-based studies for testing age-infection
401 questions like these. Meanwhile, the decreasing *F. hepatica* and *E. cervi* counts could be indicative of
402 acquired immunity over the lifespan, where older individuals become gradually more resistant due to
403 repeated exposure. This agrees with conventional wisdom in livestock that many ungulates can acquire
404 an element of immunity to *F. hepatica* infection [57], but disagrees with observations of increased *F.*
405 *hepatica* prevalence in older age categories taken from wild studies [58].

406 It is unclear how and why age-related trends would diverge for strongyles compared to *F. hepatica* and
407 *E. cervi*, and why acquired immunity might play a greater role for the latter two rather than the former.
408 Confirming a role for immunity would require 1) measuring a suite of immune traits to examine how
409 they change with age, and 2) examining whether they correlate with parasites and could therefore
410 represent immune resistance (i.e., the ability to reduce parasite load) [59]. Given that the strongyle
411 counts were measured at the order level, and generally comprise a mixture of different species, one
412 possibility is that even within this parasite count there is age-related change in the community, with
413 certain species dominating in early years that are then replaced by higher-intensity infections with other
414 species. Related, coinfecting parasites could interact with each other, either facilitating or preventing
415 each other establishing an infection in ways that contribute to the age-related changes we observe
416 [60]. For example, if strongyles and tissue worms compete indirectly by invoking the same immune
417 responses, age-related increases in strongyle intensity could result in a concurrent decrease in tissue
418 worm count. Confirming community-level changes like these would require more precise taxonomic
419 identification of the constituent nematodes, e.g. through DNA-based approaches [61,62]. A similar
420 trend is less likely for the fluke and tissue worm counts, as these are more likely to be counts of single
421 homogenous species. Ultimately, the fact that these reputedly-similar macroparasites showed highly
422 divergent trends with age is interesting, and invites further investigation.

423 Overall, our results confirm that age-related changes in infection can vary substantially within the same
424 system, and likely depend on a complex combination of immune, behavioural, and demographic
425 processes. Although we did not test specific immunological drivers of the trends we observed, this study
426 suggests that changes in exposure and demography through the lifespan could play a complex role in
427 age-infection interrelationships, and that changes in intrinsic (i.e. physiological) traits might be relatively
428 more important. Given the highly divergent age trajectories observed, this study confirms that ageing
429 individuals may not necessarily experience a greater overall parasite burden, but a different parasite
430 community, which may exert complex pressures on the age structure of the population. Understanding
431 how and why parasite community structure changes with host age – and the relative role of
432 susceptibility and exposure in determining it – is likely to provide new insight into disease transmission
433 and the ageing process in natural systems.

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