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# Oil and coal price shocks and coal industry returns: international evidence

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## Abstract

This paper examines the effect of energy price shocks on coal sector stock returns and supplements studies evaluating the effect of oil prices on the stock price of oil and gas companies. A 1% increase in coal price return raises coal sector returns by between 0.22% and 0.30%. This result is robust across developed, emerging and differing groups of Asia-Pacific and Pacific countries, and is analogous with findings that a 1% increase in oil price raises the return of oil and gas companies by between 0.14% and 0.38% depending on country and time period studied. Oil price return also significantly influences coal sector return even controlling for coal price return. Relatively large increases in coal and oil price returns have statistically significant and disproportionate effects on raising coal sector returns. Market return, interest rate premium, and foreign exchange rate risk are also significant risk factors for excess coal sector stock returns. The sensitivity of coal sector returns to oil price shocks suggest a role for investment in stocks that rise when energy prices increase in a well balanced portfolio and in pursuing profitable investment strategies. Natural gas price returns do not influence coal sector returns in the presence of coal price returns.

JEL Classifications: G12, G15, Q4

Keywords: oil price shocks, coal price shocks, coal stock price

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## **Oil price shocks and coal industry returns: international evidence**

### **1. Introduction**

The connection between oil price and stock market returns has been examined in the literature with somewhat mixed results. In an early paper, Chen *et al.* (1986) finds that for the most part oil prices do not influence stocks prices. Jones and Kaul (1996) in an investigation of the effect of oil prices on stock returns in Canada, Japan, U.K. and U.S., establish a link through changes in cash flows on stock prices in Canada and the U.S. Sadorsky (1999) finds a negative relationship between oil price shocks and aggregate stock returns for the U.S. In contrast to Huang *et al.* (1996) who find no significant effect, Ciner (2001) finds a negative connection between real stock returns and oil price futures. Recent work reporting that oil price increases lead to reduced stock returns includes O'Neil *et al.* (2008) for the U.S., the U.K. and France, Park and Ratti (2008) for the U.S. and 12 European oil importing countries, and Nandha and Faff (2008) for global industry indices (except for extractive industries). Driesprong *et al.* (2008) find that oil price changes influence future company earnings and also discount rates. Apergis and Miller (2009) however, do not find a large effect of structural oil market shocks on stock price in eight developed countries. Malik and Ewing (2009) and Arouri *et al.* (2011) find significant volatility interaction between oil and stock market sectors.

The literature examining the effect of oil price on stock prices has paid particularly close attention to the effect on the stock prices of oil and gas companies. Sadorsky (2001) and Boyer and Filion (2007) find that positive oil price shocks significantly raise stocks returns for Canadian oil and gas companies, and El-Sharif *et al.* (2005) and Mohanty and Nandha (2011) find a similar result for U.K. and U.S. oil and gas companies, respectively. Dayanandan and Donker (2011) report that oil price increases have a positive and statistically significant impact on the accounting profits of oil and gas companies in North America.

Ramos and Veiga (2011) analyse the returns of the oil and gas sector in 34 countries and find that sector returns largely depend on market portfolio and oil price returns. With regard to quantitative impact, Sadorsky (2001), for example, finds that a 1% increase in oil price raises the oil and gas equity index by 0.305%. Thus, a rise in oil price has a significant positive effect on the stock prices of oil and gas companies that is distinct from the effect of oil price on general stock price indices.

In contrast to work identifying the risk factors of the oil and gas sector and evaluating the effect of energy prices on the stock returns of oil and gas companies, relatively little similar work has appeared on the coal sector despite the importance of coal as a source of energy. Coal meets a major share of world energy requirements and is likely to continue to do so for an extended time into the future. In recent years coal provides over 23 percent of global primary energy needs (compared to 36 percent for oil), fuels 39 percent of the world's electricity industry, and provides almost 70 percent of the energy for global steel production (Statistical Review of World Energy (2009)).<sup>1</sup>

This paper examines the effect of oil shocks and coal price shocks on coal sector stock returns and will supplement studies evaluating the effect of oil prices on the stock price of oil and gas companies. We examine panel data on coal sector stock price indices available at country level and evaluate risk factors significant in determining return in the coal sector. A 1% increase in coal price return raises coal sector returns by between 0.22% and 0.30%. These results are robust across developed, emerging and groups of Asia-Pacific and Pacific countries.

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<sup>1</sup> It is also interesting that there is not much work on the connection between oil price and coal price, in contrast to papers that have investigated the connections between oil price and natural gas price, for example. Pindyck (2004) reports that crude oil price returns predict natural gas price returns (but not the other way around). Ibrahim (2009) finds that over the longer term, natural gas price adjustments to change in crude oil price. Brown and Yucel (2007) find that natural gas prices adjust to crude oil prices with such consistency that this has led to the use of rules of thumb in energy industry that relate natural gas prices to those for crude oil. An exception is work by Mohammadi (2011) who finds that crude oil prices are not influenced by coal prices and vice versa in contrast to uni-directional long-run causality from crude oil prices to natural gas prices. Radchenko (2005) reports that changes in gasoline prices lag changes in crude oil prices.

Oil price return also significantly influences coal sector return even controlling for coal price return. Natural gas prices do not influence coal sector returns in the presence of coal price returns. Oil price might influence coal sector returns since news about energy commodities focuses primarily on oil price. Research supports the view that the market for crude oil is an international market, and market participants may perceive oil price as providing information for future global demand for energy. Relatively large increases in coal and oil price returns have statistically significant and disproportionate effects on raising coal sector returns. The sensitivity of coal sector returns to oil price shocks suggest a role for investment in stocks that rise when energy prices increase in a well balanced portfolio and in pursuing profitable investment strategies.

Market return, interest rate premium, foreign exchange rate risk, and coal price returns are statistically significant in determining the excess coal sector stock returns. A multifactor market model is used to estimate the expected excess returns to coal company stock prices. Currency depreciation has a negative impact on the return of coal companies, a result similar to that found by comparable country studies for oil and gas companies. Understanding the variables that affect the behaviour of stock prices of coal companies is of importance to market participants and to policy makers, and be useful to investors and policy managers for developing efficient hedging policies for dealing with oil and energy price shocks.

The remainder of the paper is organized as follows. Section 2 describes the data and the methodology. Section 3 discusses the regression equations and oil price variables. Section 4 presents the results of the research and section 5 concludes the study.

## **2. Data and methodology**

We obtain monthly returns for coal sector indices based on the Datastream industry classification, created by FTSE and Dow Jones. The system breaks down a country's stock

market into six levels, from aggregate market level to sub-sector levels. Coal sector indices are in level 4 under the broad classification of basic resources. We find 17 indices of coal sector available at country level for Australia, Canada, Chile, China, Hong Kong, India, Indonesia, Japan, New Zealand, Poland, Philippines, Russia, Singapore, Spain, Thailand, U.K., and U.S. Data are monthly and range from January 1999 to December 2010, comprising 144 monthly observations. The excess return series for coal sector is given by natural log difference of current month's closing price from previous month's closing price minus the monthly return on short run government bond for the corresponding country. Return data are converted to U.S. dollar returns to ensure conformity of the return data across countries. Data on all variables are from Datastream.

## 2.1 Methodology

An arbitrage pricing theory approach is taken to investigate the interaction between stock returns and energy prices. To identify important determinants of coal industry stock returns we apply a multi-factor arbitrage pricing theory model to panel data. The following international factor model will be used to link priced risk factors to required rates of return in assets in the coal sector:

$$r_{i,t} = \alpha_i + \sum_{j=1}^k \beta_j F_{j,i,t} + \varepsilon_{i,t}, \quad i = 1, 2, \dots, l, \quad (1)$$

where  $r_{i,t}$  represents the excess return of the coal sector of country  $i$  at time  $t$ ,  $\beta_j$  is the factor loading or systematic risk for risk factor  $j$ , and  $F_{j,i,t}$  is the risk factor  $j$ , for country  $i$  at time  $t$ .

The variable  $\varepsilon_{i,t}$  is a random error term.  $k$  is the number of risk factors and  $l$  is the number of countries. The model is estimated assuming fixed effects using ordinary least squares and random effects panels using generalized least squares (GLS) method. Hausman test results are obtained for all specifications with the null hypothesis of no correlation between country effects and the explanatory variables (i.e. the random effects model is the null hypothesis).

## 2.2 The risk factors

In this paper we will estimate versions of equation (1) with various risk factors. In the basic model the risk factors are taken to be market return, the foreign exchange return, an interest rate differential, and coal and oil price returns. These variables affect future investment opportunities and consumption and are perceived as key variables in intertemporal asset-pricing models. The roles of market, foreign exchange rate, interest rate, and oil price as risk factors in explaining gas and oil sector returns have been examined by Sadorsky (2001), Boyer and Filion (2007), El-Sharif *et al.* ((2005), and Ramos and Viega (2011). Extensions of the basic model will consider additional risk factors that influence excess returns in the coal sector including non-linear transformations of the energy prices and measures of uncertainty about energy prices.

Global stock return and a benchmark market return of each country are used alternatively as measures of market exposure of coal sector returns. Koller *et al.* (2010) suggest that using global market index to measure market exposure of sector returns avoids possible distorted results due to the lack of diversification of the stock markets of some countries. Global stock market and local stock market indices are taken from Datastream. The excess return series for each market index, converted to U.S. dollar returns, is given by natural log difference of current month's closing price from previous month's closing price minus the monthly return on short run government bond for the corresponding country.

A short term interest rate differential is utilized as a risk factor. The interest rate differential is defined as the three-month government bond for each country and the three-month U.S. Treasury bill rate. The interest rate differential can capture differences in macroeconomic conditions and in liquidity between the countries. A higher interest rate differential indicates a less liquid monetary environment. Foreign exchange risk is measured by the monthly logarithmic difference of the U.S. dollar price of foreign currency. A fall in

the foreign exchange variable indicates a devaluation of the local currency against U.S. dollar. Data on interest rates and foreign exchange rates are from Datastream.

These variables are likely to affect coal stock price by influencing firms' expected cash flows and the discount rate at which these cash flows are discounted. The coal sector is heavily involved in international trade. The value of the local currency has an impact on revenues in the coal sector, and this in turn influences profitability and cash flow in the sector. The coal sector is capital intensive and the interest rate is an important variable in affecting return. When the interest rate fluctuates, profitability, cash flow and returns in the coal sector are affected.

The price of oil is the West Texas Intermediate (WTI) crude oil futures price contract. Sadorsky (2011) notes that the WTI crude oil futures price contract are the most widely traded oil futures contract and serve as a standard in the oil market. Boyer and Filion (2007) and Sadorsky (2001) favour futures price rather than spot price because spot prices are more affected by random noise and by transitory shortages and supplies. The price of coal is ICE Global Newcastle futures contract in U.S. dollar per metric tonne. This is the leading price benchmark for seaborne thermal coal in the Asia-Pacific region. Oil and coal price returns are given by the log difference in the monthly data for oil and coal prices. Data on oil and coal prices are from Datastream.

When coal prices increase, expected profit, profit margins and cash flow in the coal sector increases and stock price rises. Oil price increases signal greater demand for oil and for sources of energy that can substitute for oil and that can also meet the underlying demand for energy reflected in the rising price of oil. Oil prices can be expected to influence returns in the coal sector if they provide information for coal returns beyond that contained in coal price.

### **2.3. Summary statistics**



Tables 1 and 2 present summary statistics of coal sector returns and excess stock returns by country. In Table 1 most of countries have positive coal sector excess returns, with the exceptions being Hong Kong and Japan. Australia, China, Indonesia, and Thailand on the other hand have relatively high coal sector excess returns over 1999-2010. In Table 2, the emerging markets have relatively high excess stock returns compared to the developed stock markets. From Tables 1 and 2 it is evident that returns in the coal sector of a country are higher than the local market excess stock return. The coal sector also has higher standard deviation of returns than does the local stock market return.

The global stock market return is positive over 1999-2010. Kurtosis of stock market excess returns is more than three for all countries and the returns are (mostly) negative skewed. The coal sector returns also exhibit kurtosis of more than three in all markets except India. As evidenced by the Jarque-Bera statistics, both coal sector returns and local stock markets returns are not normally distributed. However, the models to be estimated are linear and normality is not presumed in order to obtain consistent estimates.

Table 3 presents summary statistics of foreign exchange and interest rate differences by country. Over the 1999-2010 period most of countries have higher short-term interest rate than U.S. The interest rate difference between the local three-month government bond rate and that for the U.S. is positive for thirteen countries and negative for Japan, Chile and Singapore. Over 1999-2010, on average, ten currencies appreciated against the U.S. dollar and six currencies did not. The standard deviation of foreign exchange rate change against the U.S. dollar is highest for Indonesia, New Zealand and Poland. Kurtosis of foreign exchange rate returns is more than three and these returns tend to be negatively skewed. The null hypothesis of Jarque-Bera tests is rejected implying the returns are not normally distributed.

Summary statistics on oil price returns, coal returns and natural gas price returns are provided at the bottom of Table 3. Oil price returns and coal price returns have positive mean

monthly values. Oil price returns are higher than the coal price returns by factor of about 65%. The standard deviation of oil price returns is also higher than that for coal price returns, but in the monthly data only by a proportion of about 8.8%. This is consistent with the finding by Pindyke (1999) that over 125 years of price data oil price was more volatile than coal price. Regnier (2007) notes that volatility and relative volatility of energy prices can vary over time depending on regulation, market structure, output elasticity and substitutability in use. Oil returns are negatively skewed and coal returns are positively skewed. The Jarque-Bera statistics imply that the null hypothesis that oil price returns are normally distributed is rejected and that the null hypothesis that coal price returns are normally distributed is not rejected.

Figure 1 displays coal price and oil price from January 1999 to December 2010. The energy prices do tend to track one another. Figure 1 reveals that there were upward jumps in prices from 2007 that continued until the Global financial crisis in September-October, 2008. During the Global financial crisis there were significant drops in oil and coal prices, with the drop in oil price occurring earlier than the drop in coal price. In the monthly data, oil price peaked in July 2008 and coal price peaked in September 2008. Prices started recovering in late 2009, with the recovery in oil price starting earlier than that in coal prices. Movement in prices between oil and coal will diverge depending on circumstances that impact relative inventories of coal and oil available to users. Coal price achieved a local peak in July 2004. During this period power generation companies experience low coal reserves during severe power shortages in China, the world's largest producer of coal. Figure 2 displays coal price return and oil price return from January 1999 to December 2010. Both the oil and coal price return series exhibit large swings in the monthly data. The Figure 2 suggests that the timing of these swings may not be that strongly related.

The correlation matrix of variables is provided in Table 4. Coal and oil price returns have a positive co-movement and correlation coefficient of 0.22. The highest correlation (0.66) is between local stock market excess return and global stock market excess return. These two variables will not appear simultaneously in the same regression. The foreign exchange rate return and local stock market return have correlation coefficient of 0.42, with the implication that there is positive co-movement between a strengthening currency and increasing local stock market returns. The pair wise correlations among the other variables are not high in absolute value. Coal and oil price returns have a positive co-movement and correlation coefficient of 0.22. Coal price return volatility has correlations of 0.01, -0.19 and -0.31, with coal price returns, oil price returns and oil price return volatility, respectively. Oil price return volatility has correlations of -0.04 and 0.18 with coal price returns and oil price returns, respectively. It is likely that overall, multicollinearity is not a problem in estimating linear regression models with these variables.

### 3. Arbitrage pricing regressions and oil price variables

#### 3.1. The basic regression

In the basic model the risk factors are taken to be market return, the foreign exchange return, an interest rate differential, and coal and oil price returns. The basic model is given by

$$r_{i,t} = \alpha + \beta_{wm} r_{wm,t} + \beta_{in} i_{i,t} + \beta_{fx} fx_{i,t} + \beta_c r_c + \beta_o r_{o,t} + \mu_{i,t}, \quad i = 1, 2, \dots, I, \quad (2)$$

where  $r_{i,t}$  represents the excess return of the coal sector of country  $i$  at time  $t$ ,  $r_{wm,t}$  represents the global market excess return at time  $t$ ,  $i_{i,t}$  is the interest rate difference between the short-term interest rate of country  $i$  and 3-month U.S. T-bill rate,  $fx_{i,t}$  is the foreign exchange return (log difference in U.S. dollar price local currency) of country  $i$ ,  $r_{c,t}$  is the coal price return,  $r_{o,t}$  is the oil price return,  $\alpha$  is a constant, and  $\mu_{i,t}$  is an error term. An alternative to

the basic model will substitute local market excess return ( $r_{lm,t}$ ) in equation (2) for global market excess return.

If the estimated coefficient of  $r_{o,t}$ ,  $\beta_o$ , is statistically significant in equation (2), then oil price return provides information for coal stock returns beyond that conveyed by coal price return. Rejection of the null hypothesis  $\beta_o < \beta_c$  indicates that oil price return is at least as important in explaining coal sector returns as is coal price returns.

In equation (2), the returns  $r_{i,t}$ ,  $r_{wm,t}$ ,  $r_{lm,t}$ ,  $r_{c,t}$  and  $r_{o,t}$  are expressed as U.S. dollar returns. A test of the null hypothesis that the exchange rate has no influence on local currency returns in the coal sector other than through the impacts on local currency denominated market (either global or local), coal and oil returns is provided by testing Ho:  $\beta_{wm} + \beta_{fx} + \beta_c + \beta_o = 1$  (or Ho:  $\beta_{lm} + \beta_{fx} + \beta_c + \beta_o = 1$ ). If the null hypothesis is not rejected, upon substitution, equation (2) becomes (the superscript L indicates local currency-denominated returns):

$$r_{i,t}^L = \alpha + \beta_{wm} r_{wm,t}^L + \beta_{in} i_{in,t} + \beta_c r_{c,t}^L + \beta_o r_{o,t}^L + \mu_{i,t}, \quad i = 1, 2, \dots, I, \quad (2')$$

with the foreign exchange term removed, since  $r_{z,t}^L \equiv r_{z,t} - fx_{i,t}$ ,  $z = i, wm, lm, c, o$ .

### 3.2. Energy price volatility

The volatility of energy price returns has also been considered as an influence on stock returns. Bernanke (1983) and Pindyck (1991) contend that uncertainty is detrimental to economic activity and that oil price return volatility has a negative impact on economic activity and the stock market. Veronesi (1999) presents a theoretical model linking economic uncertainty and stock market volatility, arguing that during periods of high uncertainty investors are more sensitive to news and that this increases asset price volatility. Sadorsky (1999) identifies oil price shocks and oil price volatility as playing an important role in explaining U.S. real stock returns. Park and Ratti (2008) state that increased volatility in

energy prices causes greater uncertainty about product demand and future returns on investment, and affects the present value of future dividends.

Oil and coal return volatility is measured as the moving average of the squared residuals obtained from AR(1) regressions for oil and coal price returns. The AR(1) regression equations are given by:

$$r_{o,t} = c_o + \varphi_o r_{o,t-1} + \varepsilon_{o,t} \quad (3a)$$

$$r_{c,t} = c_c + \varphi_c r_{c,t-1} + \varepsilon_{c,t} \quad (3b)$$

The measure of oil and coal price return volatility is given by the residuals from equations (3a) and (3b),  $\hat{\varepsilon}_{o,t}$  and  $\hat{\varepsilon}_{c,t}$ :

$$\sigma_{k,t} = \left[ (m+1)^{-1} \sum_{j=0}^m \hat{\varepsilon}_{k,t-j}^2 \right]^{0.5}, \quad k = o, c \quad (4)$$

with  $t = 0 \dots, n-m-1$  and  $m=4$ . Volatility in oil and coal price returns is based on innovations that are not explained by past oil and coal price changes. Volatility has been measured in this way by Gallant and Tauchen (1998).

An arbitrage pricing model that captures the effects of energy price volatility is given by:

$$r_{i,t} = \alpha + \beta_{wm} r_{wm,t} + \beta_{in} i_{i,t} + \beta_{fx} f_{i,t} + \beta_c r_c + \beta_o r_{o,t} + \beta_{coalvol} \sigma_{c,t} + \beta_{oilvol} \sigma_{o,t} + \mu_{i,t}, \quad i = 1, 2, \dots, l, \quad (5)$$

where volatility in coal and oil price returns is given by  $\sigma_{c,t}$  and  $\sigma_{o,t}$ , respectively.

### 3.3. Asymmetric effects of oil and coal price returns

Asymmetry in the effect of energy prices on coal sectors will also be examined. In the literature oil price increases have been found to have a greater influence in absolute value on the macroeconomic aggregates than have oil price decreases. This asymmetric effect has been documented by Mork (1989), Hooker (1996; 2002), Hamilton and Herrera (1999) and Balke *et al.* (2002), amongst others for the U.S., by Lee *et al.* (2001) and Zhang (2008) for Japan,

by Huang *et al.* (2005) for Canada, Japan and the U.S., by Cologni and Manera (2000) for the G-7, by Cunado and Garcia (2003) for most European countries, and by Lardic and Mignon (2008) for G7 and Europe and Euro area countries. It has long been noted that counter-inflationary monetary policy responses to oil price increases can lead to the appearance of asymmetric effects of oil price increases and decreases.

Hamilton (1988) argues that change in energy price creates sectoral imbalance, which in the presence of imperfect labour mobility results in short-run loss of output, which is reinforced by oil price increases and mitigated by oil price decreases. Asymmetric effects of oil price increases and decreases grounded in sectoral reallocations is reported as a basic finding by Jones *et al.* (2004) in their survey of the literature on the effects of oil price shocks. Edelstein and Kilian (2007) contend that a finding of asymmetry is due to not modelling the effects of tax reform on fixed investment and failure to disaggregate investment into energy and non-energy related investment.

Asymmetric effects of oil price on real activity have also been grounded in the effects of uncertainty on real activity. Ferderer (1996) reports that oil price changes affect oil price volatility and that the latter has a negative effect on the economy. Elder and Serletis (2010) find that controlling for oil price uncertainty reinforces the negative response in real output to higher oil prices and ameliorates the gain in real output in response to lower oil prices. Rahman and Serletis (2011) argue that negative and positive oil price shocks differ in their impact on the volatility of oil price changes. Asymmetry in the effect of oil price on real activity may also be due to asymmetric effects of crude oil price on energy prices at the retail level. Kaufmann and Laskowski (2005) argue an asymmetry arises between crude oil price and gasoline price due to refinery utilization and optimal inventory behaviour, and between crude oil price home heating oil due customer contracts.

Nandha and Faff (2008) do not observe an asymmetric effect of oil price returns on global sector returns. Park and Ratti (2008) find evidence of asymmetric effects of oil price increases and decreases on real stock returns for the U.S. and for Norway, but not for oil importing European countries. Arouri (2011) notes that asymmetric effects of oil price increases and decreases on sectoral stock prices may arise because sectors differ with regard to energy intensity of production, the use of energy associated with the final product, and the degree of imperfect competition and ability to pass on cost increases to consumers.

To test the asymmetric effect of oil price change on coal sector returns, positive change in energy price,  $r_{k,t}^{pos}$  ( $k = o, c$ ), is differentiated from negative changes in energy price,  $r_{k,t}^{neg}$  ( $k = o, c$ ), as follows:

$$r_{k,t}^{pos} = \max\{0, \ln(k_t) - \ln(k_{t-1})\} \quad k = o, c \quad (6a)$$

$$r_{k,t}^{neg} = \min\{0, \ln(k_t) - \ln(k_{t-1})\} \quad k = o, c \quad (6b)$$

where  $c_t$  is the monthly logarithmic change in ICE Global Newcastle futures coal price in US dollar per metric tonne,  $o_t$  is the 1-month future price of a barrel of WTI (in U.S. dollars).

The model is augmented by incorporating these asymmetric measures of oil and coal returns into the following equations:

$$r_{i,t} = \alpha + \beta_{wm} r_{wm,t} + \beta_i i_{i,t} + \beta_{fx} f x_{i,t} + \beta_c r_c + \beta_o^p r_{o,t}^{pos} + \beta_o^n r_{o,t}^{neg} + u_{i,t} \quad (7)$$

$$r_{i,t} = \alpha + \beta_{wm} r_{wm,t} + \beta_{in} i_{i,t} + \beta_{fx} f x_{i,t} + \beta_o r_o + \beta_c^p r_{c,t}^{pos} + \beta_c^n r_{c,t}^{neg} + \mu_{i,t} \quad (8)$$

Examination of asymmetric effect of coal (oil) price return on coal company stock will be based on inclusion of the oil (coal) price return in the regression equation. Increases and decreases in coal and in oil price should have positive coefficients in equations (7) and (8). The effect of oil price as a signal for overall energy demand could lead to asymmetric effects if rising oil price (and rising demand for energy) is expected to lead to greater use of coal in the future than falling oil price (and falling demand for energy) for energy is thought

to lead to decreased use of coal in future. A change in oil price as change in price of substitute for coal could also be asymmetric in effect, depending on the circumstances in which it is possible for substitution between these primary sources of energy. Equation (7) provides a test of the null hypothesis ( $H_0: \beta_k^{pos} = \beta_k^{neg}$ ,  $k = o, c$ ) that there is no difference between positive and negative shocks of either oil and coal price returns.

### 3.4. Net oil price and net coal price changes

The effect of large sustained increases in coal and oil prices will also be investigated. Net oil price increase, introduced by Hamilton (1996), is designed to capture how unsettling an unusually large increase in the price of oil is likely to be for the spending decisions of consumers and firms. It is argued by Lee *et al.* (1995) that oil price increases at a time when oil prices have been relatively stable is likely to have a larger effect than an increase in oil prices at a time when oil prices have been relatively volatile.

Following Hamilton (1996), net energy price increase,  $nkpi_t$  ( $k = o, c$ ), and by analogy net energy price decrease,  $nkpd_t$  ( $k = o, c$ ), are defined as:

$$nkpi_t = \max\{0, \ln(k_t) - \ln(\max(k_{t-1}, \dots, k_{t-12}))\} \quad (k = o, c) \quad (9a)$$

$$nkpd_t = \min\{0, \ln(k_t) - \ln(\min(k_{t-1}, \dots, k_{t-12}))\} \quad (k = o, c) \quad (9b)$$

Net energy price increase (decrease) measures the amount by which log price of energy exceeds (is below) its maximum (minimum) over the last twelve months. Coal sector returns might react more to a coal or an oil price return that takes coal or oil price to a twelve month high than a coal or an oil price increase that does not. These nonlinear transformations have been used in analysis of the macroeconomic effects of oil prices (see for instance Bernanke *et al.*, 1997; Lee and Ni, 2002). Kilian (2008) argues that net oil price increase may be a good measure of the exogenous component of oil price movement. Figure 3 displays the net oil price increase variable ( $nopi$ ). Net oil price increase takes on positive values in 2000,



2004-5 and 2007-8. Figure 4 displays the net coal price increase variable ( $ncpi$ ). Net coal price increase takes on larger positive values in 2004 and 2007.

A Model that captures the effects of net coal price increase and decrease and of net oil price increase and decrease is given by:

$$r_{i,t} = \alpha + \beta_{wm} r_{wm,t} + \beta_{in} i_{i,t} + \beta_{fx} fx_{i,t} + \beta_c r_c + \beta_o r_{o,t} + \beta_o^{nkpi} nkpi_t + \beta_o^{nkpd} nkpd_t + \mu_{i,t}, \quad k = o, c \quad (10)$$

Estimation of equation (10) provides a test of the hypothesis ( $H_0: \beta_o^{nopi} = 0$ ) that coal sector returns react more to an oil price return that takes oil price to a twelve month high than an oil price increase that does not. A test of the hypothesis that an oil price decline that takes oil price below the level seen in the previous twelve months has a differential impact on coal sector returns compared to an oil price decline that does not is provided by  $H_0: \beta_o^{nopd} = 0$ . Also, estimation of equation of (10) provides a test of the null hypothesis ( $\beta_o^{nopi} = \beta_o^{nopd}$ ) that coal sector returns do not react differently between oil price returns that take oil price to either a twelve month high or to a twelve month low. A similar examination can be made of the hypothesis that coal sector returns react differently to coal price returns that take coal price to a twelve month high than to coal price returns that take coal price below the level seen in the previous twelve months.

#### 4. Results

The international factor model equations for excess coal sector returns in section 3 are estimated as a panel. We estimate fixed effects using ordinary least squares and random effects panels using generalized least squares (GLS) method. Fixed effects method is advantageous if the country effects are correlated with the explanatory variables. Hausman test results are obtained for all specifications with the null hypothesis of no correlation (the random effects model is the null hypothesis). The test results for the equations show that the null hypothesis cannot be rejected in all cases. In what follows only results for random effect

panels are reported.<sup>2</sup> Data on coal sector, global and local market returns are winsorized at the 1<sup>st</sup> percentile and 99<sup>th</sup> percentile to deal with the outliers. It turns out that this procedure does not greatly affect results.

In Table 5, two sets of results are reported: in panel A with global stock market index return as market return; and in panel B with local benchmark stock index return as the market return. In each panel 6 regression equations are reported. Market excess return, the interest rate difference and foreign exchange return appear in all equations and coal and oil excess returns and volatilities appear in different combinations in equations in order to determine whether results obtained from estimating equations (2) and (5) in the text are robust. Estimates of equations (2) and (5) appear in columns 4 and 6, respectively, in Table 5. Since equation (5) is the most comprehensive of the equations estimated, the results in column 6 of Table 5 will be given most attention. In all regressions in Table 5 market excess return, the interest rate difference, foreign exchange return, coal price return, and oil return and oil return volatility are statistically significant. The Wald test statistic for panel data indicates the models are statistically significant.

In Table 5, the coefficient of global market index return,  $\beta_{wm}$ , in panel A and the coefficient of global market index return,  $\beta_{lw}$ , in panel B are statistically different from zero at 1% level of confidence. All these parameter estimates are less than 1, significantly so for the estimates of  $\beta_{lw}$  in columns (1) through (6) and for the estimates of  $\beta_{wm}$  in columns (5) and (6).<sup>3</sup> These results suggest that the equity of the coal sector is less volatile than market returns. Since in each column, the estimate of  $\beta_{lw}$  is less than  $\beta_{wm}$  it appears that coal sector returns are more sensitive to systematic risk in the global economy than to systematic risk in the local economy. In addition, the  $R^2$  results for regressions for coal sector returns are

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<sup>2</sup> The fixed effect results and Hausman test results are available upon request.

<sup>3</sup> For example, a one-tailed test that the market beta in panel A in column (6) is less than 1 has a t-statistic of 1.982 and a one-tailed test that the market beta in panel B in column (6) is less than 1 has a t-statistic of 4.576, and the 5% and 1% critical values for one-tailed tests are 1.658 and 2.358, respectively.

somewhat higher when market risk is measured by global market return than by local market return. It will be observed later that this pattern is most pronounced for coal returns in emerging economies. Thus, it is concluded that coal sector returns are strongly influenced by global market developments.<sup>4</sup>

The estimate of the coefficient of foreign exchange rate risk (a rise indicates an appreciation of the local currency) is positive and statistically significant at the 1% level in all regressions in Table 5. The appreciation of the local currency against the U.S. dollar generates positive coal industry returns, results similar to the findings of Sadorsky (2001), Boyer and Filion (2007), and Ramos and Veiga (2011) for oil and gas sector returns. The result is consistent with a money demand model in which domestic currency and stock returns move together over the cycle (Solnik and McLeavey (2009)). Real growth is associated with increased stock returns and a rise in money demand that causes a rise in the value of the domestic currency. A test of the null hypothesis that the exchange rate has no influence on local currency returns in the coal sector other than through the impacts on local currency denominated market, coal and oil returns ( $H_0: \beta_{wm} + \beta_{fx} + \beta_c + \beta_o = 1$ ) is not rejected in columns 4 and 6 of Table 5. Thus, the hypothesis that the true relationship determining local currency returns in the coal sector is given by equation (2') is not rejected.<sup>5</sup>

The estimate of the coefficient of the interest rate difference is negative and mostly statistically significant in Table 5. Tighter liquidity in a country tends to lower returns in the coal sector. This is consistent with monetary tightening signalling macroeconomic slowdown with a dampening future demand for energy. In addition, the coal sector is capital intensive

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<sup>4</sup> These results for coal sector returns are different from results found for oil and gas companies by Ferson and Harvey (1994) and Ramos and Veiga (2011). They find that if anything, local market return has a stronger influence on oil and gas sector returns than world market portfolio return.

<sup>5</sup> Faff and Brailsford (1999) report a similar outcome for most Australian sectors including the oil and gas sector, in that in an equation with all returns expressed in local currency the exchange is not statistically significant.

and higher interest rates increase the cost of carrying debt and of financing investment with negative implications for coal sector returns.

#### **4.1. Coal and oil price returns**

The coal price return is statistically significant at 1% level in determining excess return in the coal sector in all the regressions in Table 5. A 1% increase in coal price return raises the coal company returns by between 0.270% and 0.291%. The results are consistent with and analogous to findings that oil price returns are positively associated with the returns of oil and gas companies. Sadorsky (2001) and Boyer and Filion (2007), for example, find that a 1% increase in oil price raises the return of Canadian oil and gas companies by about 0.300%. Mohanty and Nandha (2011) report that a 1% increase in oil price raises return in the U.S. oil and gas sector by between 0.207% and 0.378% depending on time period. Ramos and Viegas (2011) report a smaller effect (about 0.144%) of oil price returns on returns in the oil and gas sector worldwide.

In the coal sector results in Table 5 oil price return is statistically significant at 1% level in determining excess return in the coal sector in all regressions. The magnitude of the effect of oil price return on coal sector return is sensitive to whether or not a coal price return variable appears in the regression. However, in regressions including oil and coal price returns, a 1% increase in oil price return raises coal sector returns by between 0.120% and 0.132%. Oil prices may have a sizeable impact on coal sector stock when coal price returns are included in the regression, because among energy commodities, crude oil gets more news coverage and attention by market participants and researchers. For example, Gogineni (2008) reports that during the years 2005 and 2006, oil prices figured in the headlines of *The Wall Street Journal* on 204 days, and a majority of the accompanying articles attributed stock price movements the previous day to oil price changes.

Participants in the energy markets may perceive oil price as being determined globally and as reflecting future global demand for energy overall more efficiently than does coal price. For this reason crude oil price developments have influence on coal sector stocks. Bachmeier and Griffin (2006) conclude from examination of five crude oils that the world oil market is a single integrated economic market, but the coal market is not, and that a primary global energy market overall is only existent in the long run. Humphreys and Welham (2000) observe that the coal industry by the 1990s had started to emerge as a global industry. Ekawan and Duchêne (2006) observe that the spot market had become much more important over time for trade in coal in the Atlantic region, with the fraction of spot market trade rising from 14% in 1983 to 80% of the total in 2003. It is noted by Ekawan *et al.* (2006) that spot markets have also become much more important for trade in coal in the Pacific region. Warell (2006) find that the market is globally integrated for coal. Li (2010) provides a review of the growth in an international market in steam coal and concludes that progress toward a fully developed spot market is well advanced. Li *et al.* (2010) find a stable long run cointegrating relationship between price series for coal in Europe and Japan that is supportive of a globally integrated market for coal.

In results not reported it is found that oil price risk orthogonal to coal price risk, obtained from the residuals of a regression of oil price return on coal price return, also significantly influences coal stock returns. The results imply that oil price return increases not reflected in coal price returns also have a positive effect on coal company stock returns.

#### **4.2. Coal and oil price return volatilities**

In Table 5 the result from estimating equation (5), in which the standard deviations of coal and oil price return volatilities appear, is reported in column 6. The coefficient of coal return volatility in column 6 in Table 5 is negative but is not statistically significant when

market risk factor is measured global stock market returns and is only statistically significant at the 10% level when market risk factor is measured by local stock market returns. Oil price return and volatility also appear in this equation. The coefficient of coal return volatility in column 2 in Table 5 is negative when oil price return and volatility do not appear in the regression. It is interesting that Ramos and Veiga (2011) find that increased oil price return volatility is associated with an increase in oil and gas sector returns. Thus, the response of coal sector returns to coal price return volatility contrasts with results observed for the response of oil and gas sector returns to oil price return volatility (when sector return is regressed solely on own product price return volatility).

Oil price return volatility has a negative statistically significant effect at the 1% level on coal sector returns. This return holds when market risk factor is measured global stock market returns (panel A) and by local stock market returns (panel B). An increase in oil price return volatility by its mean value decreases coal sector returns by 13.04% (9.93%) when market risk factor is measured by global stock market returns (local stock market returns).<sup>6</sup> This result is in line with that reported by Park and Ratti (2008) and Sadorsky (1999) that increased volatility in oil price reduces stock price returns measured by a general index.

### **4.3. Different groups of countries**

This section examines whether results are sensitive to the groups of countries considered. Issues that arise concern differing degrees of integration into world market by sectors in emerging countries and the differing effect of coal and oil price indices on coal sector returns in different markets.

#### **4.3.1. Developed countries vs. Emerging countries**

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<sup>6</sup> The mean of oil (coal) price return volatility defined in equation (4) is 0.0867 (0.0796). The product of the coefficient of oil price return volatility in Table 5, column 6, panel A (B), -1.5041 (-1.1458), and 0.0867 yields -0.1304 (-0.0993).

The issue of whether risk factors in coal sector returns differ between developed and emerging countries is investigated in this section. Emerging markets may not be fully integrated into the global economy and this may give rise to differences in the effect of the risk factors on coal sector returns. Carrieri and Majerbi (2006) report that returns in emerging markets are affected more by local than by global risk factors. Basher and Sadorsky (2006) find that stock markets of emerging countries are more exposed to oil price risk factor than stock markets in developed countries. Table 6 presents results of the GLS panel estimation of coal sector returns in developed countries in column 1 and in emerging countries in column 2. Developed and emerging markets are identified according to MSCI classification.<sup>7</sup>

The goodness of fit of the regressions measured by  $R^2$  is better for explaining coal sector returns in developed markets than in emerging markets, reflecting the greater volatility in general in returns in the emerging markets. In column 1 for developed markets it doesn't much matter whether the market risk factor is measured by a global market index or a local market index, since developed markets are well integrated into the global market. In column 2 for emerging markets coal sector returns are more exposed to global market systematic risk than to local market systematic risk. However, coal sector returns in emerging markets are less exposed to global market systematic risk than are coal sector returns in developed markets. In the regression equations disaggregated by developed and emerging markets, although the estimated coefficient of the interest rate difference is negative it is no longer statistically significant in most regressions. Foreign exchange rate risk is statistically significant at the 1% level for both the developed and emerging markets in regressions with global market risk and less so in regressions with local market risk.

The coefficients of coal price return and oil price return are positive and statistically significant in regressions for coal sector returns in both developed and emerging markets. In

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<sup>7</sup> Developed countries are Australia, Canada, Hong Kong, Japan, New Zealand, Singapore, Spain, U.K. and U.S. Emerging countries are Chile, China, India, Indonesia, Poland, Philippines, Russia and Thailand. Ramos and Veiga (2011) use MSCI classification in their study of risk factors in oil and gas industry returns.

Panel A with global market risk, the exposure of coal sector return to coal price return is greater than that to oil price return for both developed and emerging markets. This result is unchanged for the developed countries but is changed for emerging countries when local market risk is substituted for global market risk.

#### **4.3.2. Asia-Pacific and Pacific countries**

Robustness of results will now be examined for Asia-Pacific and Pacific countries. This will provide a check of robustness of results across regions where the ICE Global Newcastle futures contract coal price is the leading price benchmark for seaborne thermal coal. Four sub-groups are considered. Asia-Pacific<sup>1</sup> countries are Australia, Canada, Chile, China, Hong Kong, India, Indonesia, Japan, New Zealand, Philippines, Russia, Singapore, Thailand and U.S. Pacific<sup>1</sup> countries are Australia, Canada, Chile, China, Hong Kong, Indonesia, Japan, New Zealand, Philippines and Singapore. Asia-Pacific<sup>2</sup> countries are Asia-Pacific<sup>1</sup> countries minus Russia and the U.S. Pacific<sup>2</sup> countries are Pacific<sup>1</sup> countries minus China and Hong Kong.

Estimates of regression equation (2) are reported in columns 3 through 6 for these four groups of countries. It is found that coal sector returns in the groups of Asia-Pacific and Pacific markets are exposed to global market systematic risk, foreign exchange and interest rate risk, and coal price and oil price return. Coal and oil price return have statistically significant effects on coal sector returns across different groups of country. A test of the null hypothesis that the exchange rate has no influence on local currency returns in the coal sector other than through the impacts on local currency denominated market, coal and oil returns is not rejected in columns 3 through 6 in Table 6 for any of the country groups.

#### **4.4. Asymmetric effects of coal price and oil price changes**

Test results for an asymmetric effect of oil and coal price changes on coal sector returns are reported in Table 7. Estimates of equations (7) and (8) for positive and negative



oil and coal price returns are reported in columns 1 and 2, respectively, and estimates of equation (10) for net oil and coal price returns are reported in columns 3 and 4, respectively. Positive change in coal price,  $r_{c,t}^{pos}$ , is statistically significant at the 1% level of confidence in column (1) and positive change in oil price,  $r_{o,t}^{pos}$ , is statistically significant at the 1% level of confidence in column (2). The coefficients of the negative oil and price changes are also statistically significant in columns (1) and (2), but are smaller in magnitude than the coefficients of the positive oil and price changes. The null hypothesis that positive and negative coal price shocks have the same coefficient is rejected at the 1% level of confidence and the null hypothesis that positive and negative oil price shocks have the same coefficient is rejected at the 10% level of confidence. These results suggest that coal (oil) price increases have a larger positive impact on coal sector returns than coal (oil) price decreases have on decreases in coal sector return.

In column 3 of Table 8 the coefficient of net coal price increase is statistically significant at 5% level. The coefficient of net coal price decrease is negative in column 3. A Chi-square test of the null hypothesis  $\beta_c^{ncpi} = \beta_c^{ncpd}$  is rejected at the 1% level. In column 4 of Table 7 the coefficient of net oil price increase is statistically significant at the 1% level of confidence. A positive value for net oil price indicates that oil price is trading at a higher price than that observed over the previous twelve months. Coal sector returns react more to an oil price return that takes oil price to a twelve month high than an oil price increase that does not. The coefficient of net oil price decrease is statistically significant at the 10% level. The coefficient of  $nopi$  is larger than that of  $nopd$ . A Chi-square test of the null hypothesis  $\beta_o^{nopi} = \beta_o^{nopd}$  is rejected at the 1% level. Thus, oil price declines that take oil price below the level seen in the previous twelve months does have a larger impact than a regular oil price

decline (at least at the 10% level of confidence) but this differential effect is not as marked as that for oil prices breaking higher levels.

The pass-through effect of coal and of oil price returns for coal sector returns are similar to those observed by Ramos and Veiga (2011) for oil price returns on oil and gas sector returns, in that coal and oil price increases have larger effects than oil price decreases. In column 5 of Table 7 net oil and coal price increases and decreases appear together. The asymmetry between positive and negative net oil and coal price changes is again confirmed. Thus, it can be said that the asymmetry effect is observed in the coal sector returns.

#### **4.5. Natural gas price returns**

We augment this study by evaluating the effect of natural gas price returns on the coal sector returns. This allows examination of whether controlling for natural gas returns renders the influence of oil price returns on coal sector returns insignificant. Coal and natural gas are energy sources used for electricity and heating production and not considering the influence of gas price returns might bias results. In our work we use the log difference of monthly Henry Hub future price of natural gas- the leading price in natural gas market (a U.S. dollar index). From Table 2 it can be seen that gas price returns are slightly less than coal price returns over 1990:01 to 2010:12. The standard deviation of gas price returns is over twice that for either coal price returns or oil price returns. As for coal price returns (and not for oil price returns) the Jarque-Bera statistic implies that the null hypothesis that gas price returns are normally distributed is not rejected. In Table 4 gas and coal price returns have a positive co-movement and correlation coefficient of 0.07, and gas and oil price returns have a positive co-movement and correlation coefficient of 0.15. The values of these correlations indicate that inclusion of oil, coal and gas returns in the same regression do not raise multicollinearity issues.

We use the following model to evaluate the effect of gas price returns on coal sector returns:

$$r_{i,t} = \alpha + \beta_{wm} r_{wm,t} + \beta_{in} i_{i,t} + \beta_{fx} fx_{i,t} + \beta_c r_c + \beta_o r_{o,t} + \beta_g r_g + \mu_{i,t}, \quad (11)$$

where  $r_g$  is gas price return. Results from estimating equation (11) are reported in Table 8. When gas price return is the only energy price appearing in the regression equation, the coefficient of natural gas price is significant at 10% level (column 1). However, when oil price return appears in the regression equation the coefficient of  $r_g$  is not statistically significant (in column 2 and 3 of Table 8). Both coal and oil price returns are statistically significant in the presence a gas price return variable, with coefficients of 0.11 and 0.24, respectively in column 3 of Table 8. The null hypothesis that the effect of oil price return on coal sector return is less than that of coal price return on coal sector return ( $H_0: \beta_o < \beta_c$ ) is rejected at the 1% level. Thus, the result that oil price return has a larger impact on coal sector return than does coal price return is not affected by inclusion of gas price return in the regression equation.

## 5. Conclusion

In this paper we examine panel data on coal sector stock price indices available at country level and evaluate risk factors significant in determining return in the coal sector. The paper studies the effect of energy shocks on coal sector stock returns and supplements research evaluating the effect of oil prices on the stock price of oil and gas companies. A 1% increase in coal price return raises the coal company returns by between 0.27% and 0.29%. This result is robust across developed, emerging and differing groups of Asia-Pacific and Pacific countries. The results are consistent with analogous findings that a 1% increase in oil price raises the return of oil and gas companies by between 0.14% and 0.38% depending on country and time period studied.

The paper finds that oil prices have a significant impact on coal sector returns even in the presence of coal price returns. A 1% increase in coal price raises coal sector returns by about 0.12%. This result may follow because news about energy commodities focuses primarily on oil price. Research supports the view that the market for crude oil is an international market, whereas the market for coal is only more recently emerging as a global market. Participants in the market may perceive oil price as serving as the bench mark for future global demand for energy overall. For this reason crude oil price developments have influence on coal sector stocks. Natural gas prices do not influence coal sector returns in the presence of coal price returns.

Coal sector returns react more to an coal price return that takes coal price to a twelve month high than an coal price increase that does not. The coal sector responds more to a positive coal price change than to negative coal price change. It should be noted that estimation of asymmetric effects of coal price change does not erode the statistical significance of oil price change in affecting on coal sector returns. An asymmetry in the effect of oil prices on coal sector returns is also observed. Coal sector returns react more to an oil price increase than to an oil price decrease and more to an oil price return that takes oil price to a twelve month high than an oil price increase that does not. Increased volatility in oil price return significantly reduces coal sector return. Increased coal price volatility does not significantly affect coal sector return.

Market return, interest rate premium, foreign exchange rate risk, and coal price returns are statistically significant in determining the excess coal sector stock returns. Currency depreciation has a negative impact on the return of coal companies, a result similar to that found by comparable country studies for oil and gas companies. The exchange rate does not significantly influence local currency returns in the coal sector other than through the impacts on local currency denominated market, coal and oil returns. Understanding the

variables that affect the behaviour of stock prices of coal companies is of importance to market participants and to policy makers for developing efficient hedging policies for dealing with oil and energy price shocks.

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## Appendix

**Table A: Definition of Variables**

Variable	Symbol	Measures
Global market excess return	$r_{wm}$	Monthly logarithmic change in the global stock market index in excess of a 3 month Treasury bill rate. US dollar return.
Local market excess return	$r_{lm}$	Monthly logarithmic change in the local stock market index in excess of short term interest rate of corresponding market. US dollar return.
Foreign exchange rate	$fx$	Monthly logarithmic change in US dollar price of foreign currency.
Interest rate difference	$i$	Monthly difference between the short term interest rate of a country and three month US Treasury bill.
Oil price return	$r_o$	Monthly logarithmic change in West Texas Intermediate crude oil futures price per barrel. US dollar return.
Coal price return	$r_c$	Monthly logarithmic change in ICE Global Newcastle futures coal price in US dollar per metric tonne.
Natural gas price return	$r_g$	Monthly logarithmic change in Henry Hub natural gas future prices per million British Thermal Unit.
Oil return volatility	$\sigma_o$	Monthly volatility in oil price return
Coal Return volatility	$\sigma_c$	Monthly volatility in coal price return.
Positive oil price return	$r_o^{pos}$	Oil price return if positive, otherwise zero.
Negative oil price return	$r_o^{neg}$	Oil price return if negative, otherwise zero.
Positive coal price return	$r_c^{pos}$	Coal price return if positive, otherwise zero.
Negative coal price return	$r_c^{neg}$	Coal price return if negative, otherwise zero.
Net oil price increase	$nopi$	Log oil price minus maximum log oil price over preceding twelve months if positive, otherwise zero.
Net oil price decrease	$nopd$	Log oil price minus minimum log oil price over preceding twelve months if negative, otherwise zero.
Net coal price increase	$ncpi$	Log coal price minus maximum log coal price over preceding twelve months if positive, otherwise zero.
Net coal price decrease	$ncpd$	Log coal price minus minimum log coal price over preceding twelve months if negative, otherwise zero.

Notes: Data from DataStream

**Table 1: Summary statistics: coal sector returns**

Dependent Variable						
Country	Mean	SD	Kurtosis	Skewness	JB	p-value
Australia	0.0211	0.1006	5.6380	-0.3756	45.14	0.0000
Canada	0.0135	0.2816	7.0958	-0.7525	110.2572	0.0000
Chile	0.0158	0.2953	62.7240	6.5619	2243.18	0.0000
China	0.0242	0.1418	4.7187	-0.0746	17.8578	0.0000
Hong Kong	-0.0116	0.2844	5.3912	-1.0207	15.6516	0.0004
India	0.0080	0.0626	2.6554	-0.2052	0.0837	0.9590
Indonesia	0.0239	0.2311	6.4659	0.4132	24.3326	0.0000
Japan	-0.0005	0.1699	4.8130	0.7737	34.0764	0.0000
New Zealand	0.0015	0.0253	4.3816	0.5088	4.9077	0.0860
Philippines	0.0126	0.2321	4.7973	0.6948	30.9671	0.0000
Poland	0.0222	0.0941	2.9097	0.5020	0.6774	0.7127
Russia	0.0192	0.2342	4.9022	-0.9502	14.4594	0.0007
Singapore	0.0328	0.2068	3.8280	-0.3065	2.1229	0.3460
Spain	0.0005	0.0790	6.1979	0.8786	65.4522	0.0000
Thailand	0.0247	0.1337	5.5019	-0.7915	52.5921	0.0000
UK	0.0032	0.1766	18.1058	-1.9064	1456.35	0.0000
US	0.0137	0.1376	5.1317	-0.9507	48.9580	0.0000

Notes: Summary statistics of the coal sector monthly excess returns are reported by country over 1999:01 through 2010:12. Mean, standard deviation (SD), kurtosis, skewness, and Jarque-Bera (JB) statistics and p-values are reported in each column. Return is the first difference of the logarithm of coal sector price in U.S. dollars minus a short-term interest rate.

**Table 2: Summary statistics: market returns**

Independent Variable						
	Mean	SD	Kurtosis	Skewness	JB	p-value
$r_{wm}$	0.0022	0.0539	4.9842	-0.7852	38.4167	0.0000
$r_{lm}$						
Australia	0.0065	0.0681	5.2655	-0.7670	44.9035	0.0000
Canada	0.0084	0.0661	5.8545	-0.8932	68.0339	0.0000
Chile	0.0110	0.0589	5.2043	-0.5144	35.5056	0.0000
China	0.0160	0.0948	3.6116	-0.0321	2.2692	0.3216
Hong Kong	0.0061	0.0678	3.5371	-0.1296	2.1339	0.3440
India	0.0133	0.1041	3.8844	-0.3518	7.6628	0.0217
Indonesia	0.0111	0.2058	10.3124	0.1861	321.6554	0.0000
Japan	0.0007	0.0554	3.2144	-0.0631	0.3714	0.8305
New Zealand	0.0037	0.0644	3.8119	-0.6411	13.8175	0.0010
Philippines	0.0095	0.0640	4.7661	-0.3109	21.0353	0.0000
Poland	0.0067	0.1022	4.4188	-0.6137	21.1178	0.0000
Russia	0.0227	0.1193	4.6641	-0.4063	20.5763	0.0000
Singapore	0.0100	0.0764	4.6815	-0.2982	19.1000	0.0001
Spain	0.0011	0.0683	4.7838	-0.6516	29.2800	0.0000
Thailand	0.0079	0.0970	4.1966	-0.1289	8.9907	0.0112
U.K.	0.0003	0.0563	5.4545	-0.5253	42.7704	0.0000
U.S.	0.0003	0.0521	4.6841	-0.7494	30.4968	0.0000

Notes: This table reports summary statistics of global stock market excess return ( $r_{wm}$ ) and local stock market excess return ( $r_{lm}$ ) over 1999:01 through 2010:12. Mean, standard deviation (SD), kurtosis, skewness, and Jarque-Bera (JB) statistics and p-values are reported in each column. Return is the first difference of the logarithm of coal sector price in U.S. dollars minus a short-term interest rate.

**Table 3: Summary statistics: interest rate, foreign exchange rate, oil and coal price**

Independent variable						
$i$	Mean	SD	Kurtosis	Skewness	JB	p-value
Australia	0.0426	0.0118	2.2613	0.1430	1.0719	0.5851
Canada	0.0046	0.0043	2.5435	0.5828	2.6771	0.2622
Chile	-0.0070	0.0123	4.3338	-1.5432	19.3134	0.0001
China	0.0106	0.0093	5.7972	-1.6278	31.4732	0.0000
Hong Kong	0.0048	0.0071	7.1356	2.0403	57.6630	0.0000
India	0.0461	0.0155	2.2358	0.2096	1.2980	0.5226
Indonesia	0.0682	0.0182	3.7216	0.5119	2.6802	0.2618
Japan	-0.0037	0.0124	4.3813	-1.5632	19.9566	0.0000
New Zealand	0.0421	0.0179	1.6879	0.5996	5.3981	0.0673
Philippines	0.0308	0.0171	2.9599	-0.8736	5.2176	0.0736
Poland	0.0382	0.0131	4.6995	-0.9554	11.1710	0.0038
Russia	0.0810	0.0636	3.9828	1.3889	14.8310	0.0006
Singapore	-0.0017	0.0068	5.0415	-1.5326	23.1708	0.0000
Spain	0.0141	0.0095	4.7827	-0.1816	5.6546	0.0592
Thailand	0.0095	0.0093	2.2733	0.6078	3.4264	0.1803
U.K.	0.0135	0.0127	1.8534	0.7447	6.0352	0.0489
$fx$						
Australia	0.0030	0.0399	4.5354	-0.5065	20.3016	0.0000
Canada	0.0028	0.0262	4.9191	-0.2929	24.1578	0.0000
Chile	-0.0002	0.0328	6.2566	-0.9428	84.9653	0.0000
China	0.0019	0.0057	42.6706	5.3884	10139.3600	0.0000
Hong Kong	0.0000	0.0014	12.4506	1.5580	594.1347	0.0000
India	-0.0004	0.0180	5.7103	-0.3357	46.7798	0.0000
Indonesia	-0.0013	0.0457	6.5802	0.5966	85.4475	0.0000
Japan	0.0026	0.0293	3.1600	-0.2735	1.9494	0.3773
New Zealand	0.0025	0.0426	4.5709	-0.3185	17.2409	0.0000
Philippines	-0.0007	0.0198	6.4823	-0.8395	89.6725	0.0000
Poland	0.0009	0.0426	5.0187	-0.7462	37.8155	0.0000
Russia	-0.0036	0.0311	20.3432	-3.1144	2037.5040	0.0000
Singapore	0.0016	0.0162	4.3876	-0.1532	12.1156	0.0023
Spain	-0.0008	0.0315	3.6740	0.0465	2.7775	0.2494
Thailand	0.0013	0.0211	4.1751	-0.1307	8.6952	0.0129
U.K.	0.0004	0.0276	5.4816	0.3944	40.6838	0.0000
$r_o$	0.0107	0.0954	4.5885	-0.5885	13.8361	0.0000
$r_c$	0.0065	0.0877	3.4522	0.2508	2.7362	0.2546
$r_g$	0.0058	0.2144	4.5593	-0.0493	14.54	0.0007

Notes: Summary statistics for interest rate difference,  $i$ , foreign exchange rate return,  $fx$ , oil price return,  $r_o$ , coal price return,  $r_c$ , and natural gas price return,  $r_g$  are reported for 1999:01 through 2010:12. Mean, standard deviation (SD), kurtosis, skewness, and Jarque-Bera (JB) statistics and p-values are reported in each column. The interest rate difference is three month local government bond rate minus U.S. equivalent, foreign exchange rate is the log difference in the U.S. dollar price of the local currency, oil price return is the log difference in one month future price of WTI, coal price return is log difference in ICE Global Newcastle futures price of coal, and natural gas price is log difference in Henry Hub future prices.

**Table 4: Correlation matrix of the variables**

	World market return	Local market return	Foreign exchange rate return	Interest rate difference	Coal price return	Oil price return	Natural gas return	Coal price return volatility	Oil price return volatility
World market	1.0000								
Local market	0.6562	1.0000							
Foreign exchange	0.3599	0.4197	1.0000						
Interest rate difference	-0.0795	-0.0494	-0.0344	1.0000					
Coal return	-0.0636	-0.0777	0.0073	-0.0687	1.0000				
Oil return	-0.0734	-0.0842	-0.1214	-0.0494	0.2209	1.0000			
Natural gas return	-0.0303	-0.0297	-0.0276	-0.0211	0.0707	0.1543	1.0000		
Coal volatility	-0.0654	-0.0280	-0.0845	0.0876	0.0143	-0.1906	-0.0431	1.0000	
Oil volatility	-0.1262	-0.0861	0.0156	0.0520	-0.0426	0.1791	-0.0793	-0.3094	1.0000

**Table 5: Coal sector return equations and oil price shocks**

Variables	1	2	3	4	5	6
<b>Panel A</b>						
Constant	0.1801*** (0.0710)	0.2875*** (0.0654)	0.1851*** (0.0741)	0.1892*** (0.0721)	0.1650*** (0.0625)	0.3010*** (0.0874)
$r_{wm}$	0.8907*** (0.0910)	0.8085*** (0.1120)	0.8604*** (0.1147)	0.9012*** (0.1514)	0.8638*** (0.1547)	0.7541*** (0.1241)
$fx$	0.4987*** (0.1010)	0.5021*** (0.1741)	0.4174*** (0.1047)	0.4321*** (0.1925)	0.4574*** (0.1474)	0.6587*** (0.2001)
$i$	-0.4010** (0.1923)	-0.3737 (0.2325)	-0.3785** (0.1873)	-0.3768 (0.2910)	-0.4256*** (0.1900)	-0.2784 (0.2155)
$r_o$			0.1787*** (0.0410)	0.1256*** (0.0425)	0.1766*** (0.0587)	0.1198*** (0.0352)
$r_c$	0.2914*** (0.1214)	0.2985*** (0.0901)		0.2890*** (0.0680)		0.2875*** (0.0741)
$\sigma_c$		-1.1514* (0.6545)				-1.1210 (0.7985)
$\sigma_o$					-1.3785** (0.6555)	-1.5041*** (0.4123)
Wald $\chi^2$	136.20	147.59	122.54	140.36	137.10	162.14
Prob> $\chi^2$	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
$R^2$	0.1514	0.1498	0.1810	0.1817	0.1893	0.1987
$\chi^2$ test $\beta_{wm} + \beta_{fx} + \beta_c + \beta_o = 1$				4.01 (0.405)		5.19 (0.182)
<b>Panel B</b>						
Constant	0.1792*** (0.0681)	0.3020*** (0.1110)	0.1839*** (0.0741)	0.1910*** (0.0620)	0.1689*** (0.0654)	0.3008*** (0.1101)
$r_{lm}$	0.5751*** (0.0912)	0.5241*** (0.1019)	0.5325*** (0.1024)	0.5541*** (0.0743)	0.5311*** (0.0874)	0.5125*** (0.1120)
$fx_i$	0.4014*** (0.1899)	0.4820*** (0.1354)	0.4125*** (0.1641)	0.4597*** (0.1785)	0.4546*** (0.2101)	0.4987*** (0.0541)
$i_i$	-0.4641** (0.2414)	-0.3990* (0.2375)	-0.4262** (0.2120)	-0.4049 (0.2260)	-0.4594** (0.2263)	-0.2987* (0.1767)
$r_o$			0.2042*** (0.0624)	0.1321*** (0.0510)	0.2033*** (0.0347)	0.1241*** (0.0424)
$r_c$	0.2781*** (0.0674)	0.2701*** (0.0489)		0.2872*** (0.0629)		0.2698*** (0.0652)
$\sigma_c$		-1.3990* (0.8278)				-1.6988* (0.9876)
$\sigma_o$					-0.9896*** (0.3993)	-1.1458*** (0.3874)
Wald $\chi^2$	101.25	137.41	101.21	117.20	108.32	123.56
Prob> $\chi^2$	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
$R^2$	0.1152	0.1411	0.1312	0.1142	0.1312	0.1614
$\chi^2$ test $\beta_{lm} + \beta_{fx} + \beta_c + \beta_o = 1$				5.21 (0.266)		7.24 (0.123)

Notes: This table reports results from estimating versions of equation (5):

$$r_{i,t} = \alpha + \beta_{wm}r_{wm,t} + \beta_{in}i_{i,t} + \beta_{fx}fx_{i,t} + \beta_c r_c + \beta_o r_{o,t} + \beta_{coalvol}\sigma_{c,t} + \beta_{oilvol}\sigma_{o,t} + \mu_{i,t}$$

The dependent variable is the monthly excess returns of the coal industry indices in U.S. dollars. Explanatory variables include the global market return ( $r_{wm}$ ) or local market return ( $r_{lm}$ ), the log difference in the U.S. dollar price of local currency ( $fx$ ), difference between the local interest rate and the U.S. interest rate ( $i$ ), coal price return ( $r_c$ ), oil price return ( $r_o$ ), volatility of coal returns ( $\sigma_c$ ), and volatility of oil returns ( $\sigma_o$ ). The model is estimated using random effects GLS method, since the null hypothesis of no correlation of country effects being correlated with the explanatory variables is not rejected. The standard errors robust to heteroskedasticity appear in parentheses below parameter estimates, and errors are clustered by country. P-value appears below  $\chi^2$  test results. Statistical significance at 1%, 5% and 10% levels of confidence is indicated by \*\*\*, \*\* and \*, respectively.



**Table 6: Coal sector return equations for different groups of countries**

Variables	1	2	3	4	5	6
Panel A	Developed	Emerging	Asia-Pacific <sup>1</sup>	Asia-Pacific <sup>2</sup>	Pacific <sup>1</sup>	Pacific <sup>2</sup>
Constant	0.1582** (0.0815)	0.1741*** (0.0741)	0.1498** (0.0752)	0.1751*** (0.0551)	0.1513*** (0.0452)	0.1684*** (0.0447)
$r_{wm}$	1.1001*** (0.1910)	0.7354*** (0.1474)	1.0432*** (0.1785)	0.8874*** (0.1891)	0.9452*** (0.1525)	0.9573*** (0.2150)
$f_x$	0.4871*** (0.2010)	0.5474*** (0.2010)	0.4258*** (0.1987)	0.4987*** (0.1874)	0.3952*** (0.2014)	0.4235** (0.2090)
$i$	-0.1618 (0.1241)	-0.0154 (0.1024)	-0.1941* (0.1132)	-0.0987 (0.0856)	-0.2014 (0.1293)	-0.1118 (0.0987)
$r_o$	0.0612** (0.0309)	0.0754*** (0.0310)	0.0834** (0.0380)	0.0971*** (0.0298)	0.0925*** (0.0350)	0.1025*** (0.0289)
$r_c$	0.2219*** (0.0914)	0.2651*** (0.0698)	0.2180*** (0.0825)	0.2515*** (0.0791)	0.2421*** (0.0920)	0.2987*** (0.0474)
Wald $\chi^2$	214.10	114.37	224.10	184.21	190.20	175.21
Prob> $\chi^2$	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
$R^2$	0.1587	0.1021	0.1710	0.1874	0.1982	0.1692
$\chi^2$ test: $\beta_{wm} + \beta_{fx} + \beta_c + \beta_o = 1$						
	5.96 (0.114)	5.56 (0.135)	6.20 (0.102)	4.01 (0.260)	4.89 (0.180)	5.19 (0.158)
Panel B						
Constant	0.2014** (0.0921)	0.1914** (0.0952)	0.1479** (0.0752)	0.1821*** (0.0624)	0.1415*** (0.0474)	0.1897*** (0.0503)
$r_{lm}$	0.9874*** (0.1541)	0.4825*** (0.0741)	0.6051*** (0.1025)	0.5959*** (0.1751)	0.5941*** (0.1012)	0.5785*** (0.1954)
$f_{x_i}$	0.2587* (0.1478)	0.3687** (0.1756)	0.4021** (0.1975)	0.4874*** (0.1984)	0.3852*** (0.2062)	0.4354** (0.2117)
$i_i$	-0.3021*** (0.1124)	-0.0541 (0.1974)	-0.2052* (0.1078)	-0.1025 (0.0765)	-0.2214* (0.1285)	-0.1285 (0.0887)
$r_o$	0.0874** (0.0470)	0.1895*** (0.0410)	0.0874*** (0.0299)	0.0920*** (0.0301)	0.0895*** (0.0350)	0.1014*** (0.0251)
$r_c$	0.2203*** (0.0889)	0.1474*** (0.0477)	0.2211*** (0.0901)	0.2458*** (0.0758)	0.2335*** (0.0852)	0.2884*** (0.0521)
Wald $\chi^2$	184.10	115.17	152.10	161.21	159.20	165.21
Prob> $\chi^2$	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
$R^2$	0.1610	0.1008	0.1524	0.1628	0.1658	0.1705
$\chi^2$ test: $\beta_{lm} + \beta_{fx} + \beta_c + \beta_o = 1$						
	3.98 (0.263)	3.12 (0.373)	4.72 (0.194)	5.21 (0.157)	4.57 (0.261)	6.24 (0.110)

Notes: Table 6 reports results from estimating version equation (2):

$$r_{i,t} = \alpha + \beta_{wm}r_{wm,t} + \beta_{in}i_{i,t} + \beta_{fx}fx_{i,t} + \beta_c r_c + \beta_o r_{o,t} + \mu_{i,t}$$

The dependent variable is the monthly excess returns of the coal industry indices in U.S. dollars. Explanatory variables include the global market return ( $r_{wm}$ ) or local market return ( $r_{lm}$ ), the log difference in the U.S. dollar price of local currency ( $fx$ ), difference between the local interest rate and the U.S. interest rate ( $i$ ), coal price return ( $r_c$ ), oil price return ( $r_o$ ), volatility of coal returns ( $\sigma_c$ ), and volatility of oil returns ( $\sigma_o$ ).

Country groups are the following. Developed countries are Australia, Canada, Hong Kong, Japan, New Zealand, Singapore, Spain, U.K. and U.S. Emerging countries are Chile, China, India, Indonesia, Poland, Philippines, Russia and Thailand. Asia-Pacific<sup>1</sup> countries are Australia, Canada, Chile, China, Hong Kong, India, Indonesia, Japan, New Zealand, Philippines, Russia, Singapore, Thailand and U.S. Asia-Pacific<sup>2</sup> countries are Asia-Pacific<sup>1</sup> countries minus Russia and the U.S. Pacific<sup>1</sup> countries are Australia, Canada, Chile, China, Hong Kong, Indonesia, Japan, New Zealand, Philippines and Singapore. Pacific<sup>2</sup> countries are Pacific<sup>1</sup> countries minus China and Hong Kong.

The model is estimated using random effects GLS method, since the null hypothesis of no correlation of country effects being correlated with the explanatory variables is not rejected. The standard errors robust to heteroskedasticity appear in parentheses below parameter estimates, and errors are clustered by country. P-value appears below  $\chi^2$  test results. Statistical significance at 1%, 5% and 10% levels of confidence is indicated by \*\*\*, \*\* and \*, respectively.

**Table 7: Asymmetric effects of coal price shocks**

Variables	1	2	3	4	5
Constant	0.1842*** (0.0662)	0.1742** (0.0762)	0.1407* (0.0762)	0.2141*** (0.0829)	0.1595*** (0.0537)
$r_{wm}$	0.9547*** (0.0942)	0.9969*** (0.0974)	0.9819*** (0.0962)	0.9224*** (0.1987)	1.1045*** (0.1199)
$fx$	0.4839** (0.2099)	0.4839** (0.2099)	0.4563*** (0.1934)	0.5958*** (0.1876)	0.4413** (0.2205)
$i$	-0.1952** (0.1033)	-0.4304*** (0.1352)	-0.4074*** (0.1787)	-0.1587 (0.1110)	-0.2219*** (0.0795)
$r_o$	0.1121*** (0.0321)		0.1249*** (0.0413)	0.1274*** (0.0587)	0.1029*** (0.0391)
$r_c$		0.1748*** (0.0513)	0.2144*** (0.0601)	0.2587*** (0.0740)	0.2137*** (0.0528)
$r_c^{pos}$	0.3237*** (0.1446)				
$r_c^{neg}$	0.1317*** (0.698)				
$r_o^{pos}$		0.1252*** (0.0421)			
$r_o^{neg}$		-0.0801* (0.0493)			
$ncpi$			0.1125** (0.0559)		0.0941** (0.0437)
$ncpd$			-0.0352* (0.0194)		-0.0122 (0.0130)
$nopi$				0.1010*** (0.0254)	0.0997** (0.0502)
$nopd$				0.0641* (0.0377)	0.0536** (0.0269)
$\chi^2$ test $\beta_c^p = \beta_c^n$	14.22** (0.0213)				
$\chi^2$ test $\beta_o^p = \beta_o^n$		5.22* (0.0750)			
$\chi^2$ test $\beta_c^{ncpi} = \beta_c^{ncpd}$			7.55** (0.0229)		5.13* (0.0769)
$\chi^2$ test $\beta_o^{nopi} = \beta_o^{nopd}$				19.25*** (0.0001)	11.56*** (0.0001)
Wald $\chi^2$ Prob> $\chi^2$	213.22 (0.0000)	195.51 (0.0000)	206.89 (0.0000)	185.14 (0.0000)	
$R^2$	0.1821	0.1372	0.1526	0.1241	

Notes: This table reports results from estimating versions of equations (7), (8) and (10):

$$r_{i,t} = \alpha + \beta_{wm} r_{wm,t} + \beta_{in} i_{i,t} + \beta_{fx} fx_{i,t} + \beta_c r_c + \beta_o^p r_{o,t}^{pos} + \beta_o^n r_{o,t}^{neg} + \mu_{i,t}$$

$$r_{i,t} = \alpha + \beta_{wm} r_{wm,t} + \beta_{in} i_{i,t} + \beta_{fx} fx_{i,t} + \beta_o r_o + \beta_c^p r_{c,t}^{pos} + \beta_c^n r_{c,t}^{neg} + \mu_{i,t}$$

$$r_{i,t} = \alpha + \beta_{wm} r_{wm,t} + \beta_{in} i_{i,t} + \beta_{fx} fx_{i,t} + \beta_c r_c + \beta_o r_{o,t} + \beta_o^{ncpi} nkpi_t + \beta_o^{ncpd} ncpd_t + \mu_{i,t}, \quad k = o, c$$

The dependent variable is the monthly excess returns of the coal industry indices in U.S. dollars. Explanatory variables are global market return ( $r_{wm}$ ), the log difference in the U.S. dollar price of local currency ( $fx$ ), difference between the local interest rate and the U.S. interest rate ( $i$ ), oil price return ( $r_o$ ), coal price return ( $r_c$ ), positive oil price returns ( $r_o^{pos}$ ), negative oil price returns ( $r_o^{neg}$ ), positive coal price returns ( $r_c^{pos}$ ), negative coal price returns ( $r_c^{neg}$ ), net oil price increase ( $nopi$ ), net oil price decrease ( $nopd$ ), net coal price increase ( $ncpi$ ), and net coal price decrease ( $ncpd$ ). The model is estimated using random effects GLS method,

since the null hypothesis of no correlation of country effects being correlated with the explanatory variables is not rejected. The standard errors robust to heteroskedasticity appear in parentheses below parameter estimates, and errors are clustered by country. P-value appears below  $\chi^2$  test results. Statistical significance at 1%, 5% and 10% levels of confidence is indicated by \*\*\*, \*\* and \*, respectively.

**Table 8: Effect of natural gas price returns**

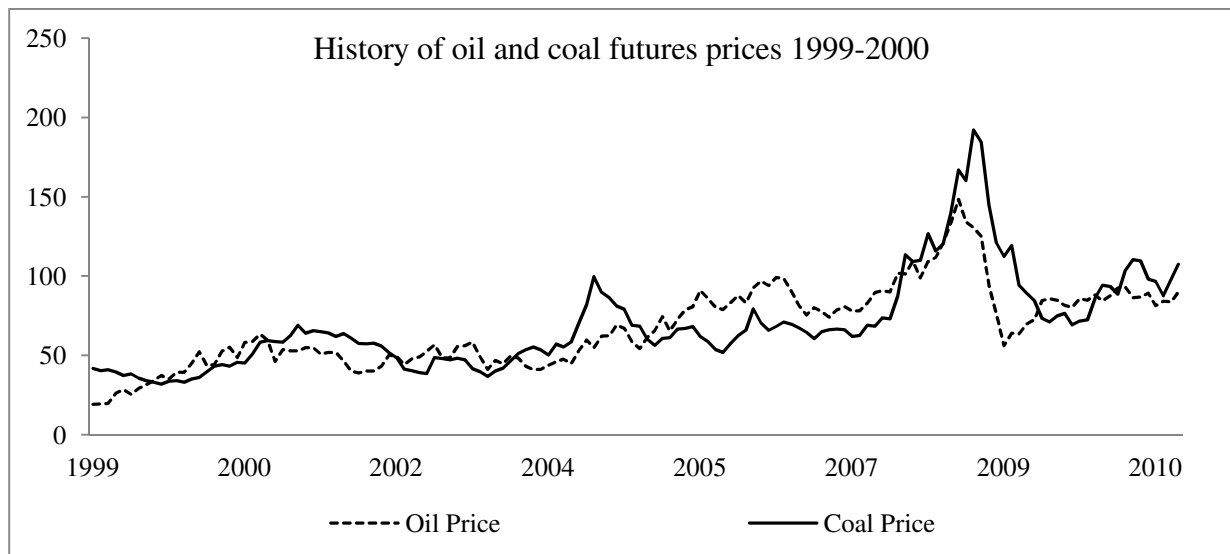
Variables	1	2	3
Constant	0.1028*** (0.0325)	0.1713*** (0.0458)	0.1598*** (0.0421)
$r_{wm}$	0.8862*** (0.0915)	0.8729*** (0.0743)	0.8652*** (0.0852)
$fx$	0.3921*** (0.1421)	0.3809*** (0.1400)	0.3658*** (0.0741)
$i$	-0.1915*** (0.0742)	-0.0499* (0.0265)	-0.0645* (0.0361)
$r_o$			0.1093*** (0.0407)
$r_c$		0.2610*** (0.0611)	0.2377*** (0.0419)
$r_g$	0.0403* (0.0232)	0.0419 (0.0311)	0.0355 (0.0287)
Wald $\chi^2$	284.12	296.39	342.90
Prob> $\chi^2$	0.0000	0.0000	0.0000
$R^2$	0.1720	0.1854	0.2317

Notes: This table reports estimation results from following equation (12)

$$r_{i,t} = \alpha + \beta_{wm} r_{wm,t} + \beta_{in} i_{i,t} + \beta_{fx} fx_{i,t} + \beta_c r_c + \beta_o r_o + \beta_g r_g + \mu_{i,t}$$

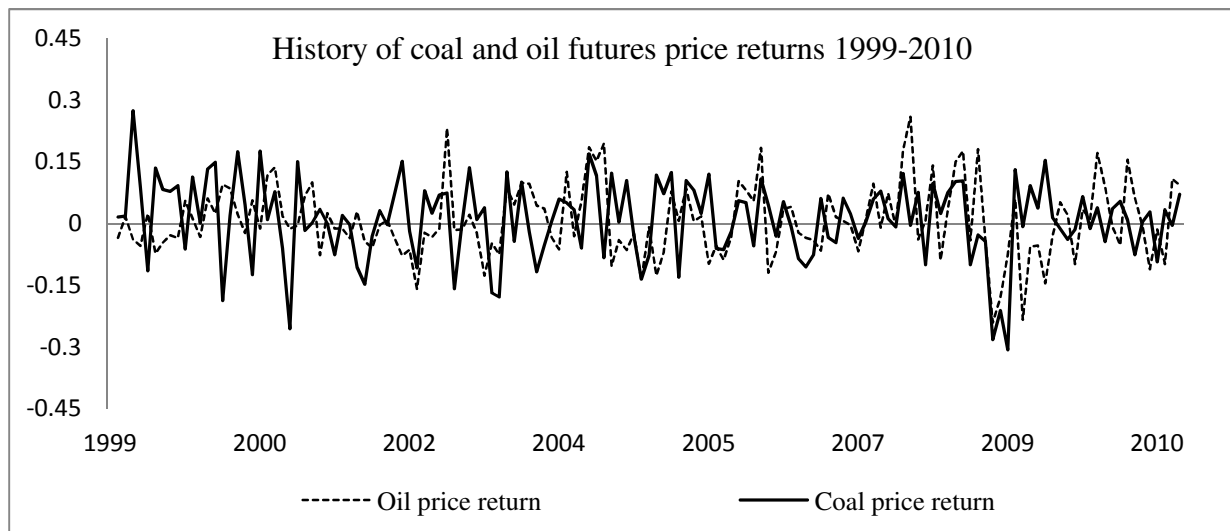
Explanatory variables are global market return ( $r_{wm}$ ), the log difference in the US dollar price of local currency ( $fx$ ), difference between the local interest rate and the US interest rate ( $i$ ), oil price return ( $r_o$ ), coal price return ( $r_c$ ), natural gas price return ( $r_g$ ). The model is estimated using random effects GLS method, since the null hypothesis of no correlation of country effects being correlated with the explanatory variables is not rejected. The standard errors robust to heteroskedasticity appear in parentheses below parameter estimates, and errors are clustered by country. P-value appears below  $\chi^2$  test results. Statistical significance at 1%, 5% and 10% levels of confidence is indicated by \*\*\*, \*\* and \*, respectively.

**Figure 1: Oil and coal futures prices in US dollars**



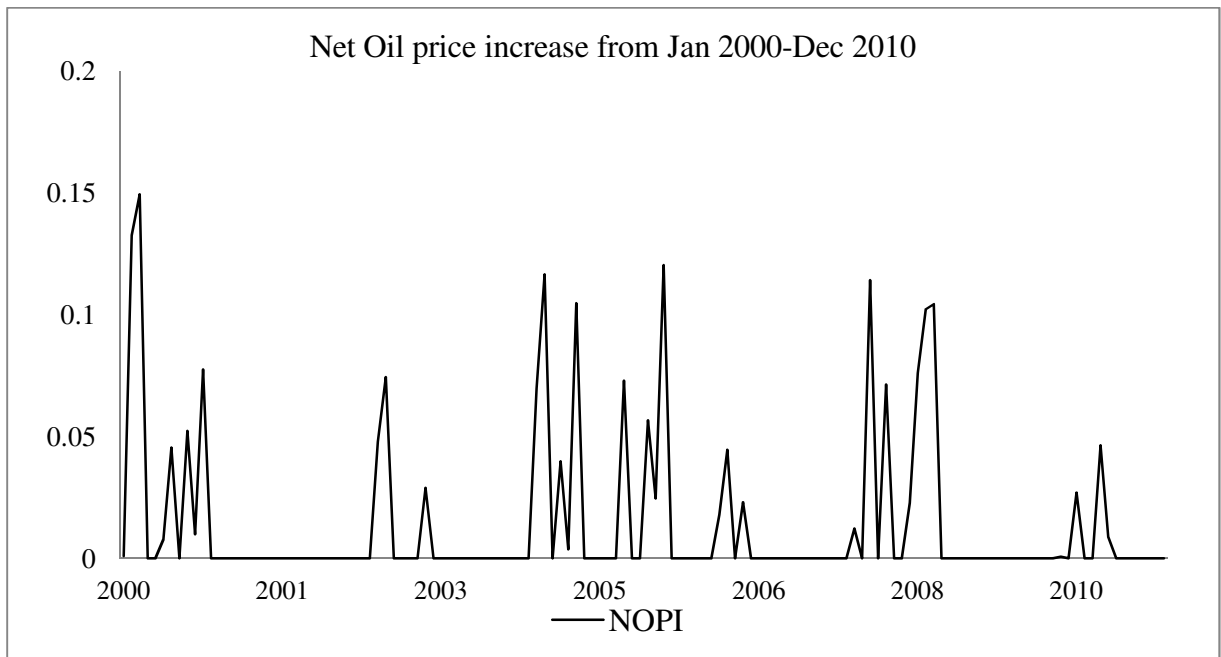
Notes: Oil price is monthly West Texas Intermediate crude oil futures price in US dollars per barrel. Coal price return is monthly ICE Global Newcastle futures coal price in US dollar per metric tonne. Data are from Datastream.

**Figure 2: Oil and coal futures price returns**



Notes: Oil price return is monthly logarithmic change in West Texas Intermediate crude oil futures price in US dollars per barrel. Coal price return is monthly logarithmic change in ICE Global Newcastle futures coal price in US dollar per metric tonne. Data are from Datastream.

**Figure 3: Net oil price increase (NOPI)**



**Figure 4: Net coal price increase (NCPI)**

