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4 **The Role of Loneliness and Social Isolation in Mediating**  
5 **the Relationship Between Childhood Maltreatment and Schizophrenia:**  
6 **A Genetically Informed Approach**  
7

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10

11 **Author Note**

12 VB, GZ and EW conceived the study, GZ compiled the data and GZ and VB  
13 conducted the analyses. LHW provided critical feedback on the draft. All authors drafted and  
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19 EP/Y015037/1]. We have no conflict of interest to disclose. The code behind this analysis has  
20 been made publicly available on GitHub at [https://github.com/VilteBaltra/loneliness-](https://github.com/VilteBaltra/loneliness-mediation)  
21 [mediation](https://github.com/VilteBaltra/loneliness-mediation). Some of the results in the manuscript have previously been presented in  
22 preliminary form at the European Neuropsychopharmacology conference (2022).

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### Abstract

Observational studies have found loneliness and social isolation to mediate the relationship between childhood maltreatment and schizophrenia. Limitations with observational studies (e.g., confounding and reverse causation), however, have meant the robustness of these relationships has thus far not been explored. To address this gap, the current study utilised genomic structural equation modelling (Genomic SEM) and Mendelian randomisation (MR) to perform a genetic mediation analysis between childhood maltreatment, loneliness/isolation, and schizophrenia, using summary statistics from three genome-wide association studies (sample sizes 105,318-487,647). Whilst we observed a putative effect of both childhood maltreatment (inverse variance weighted [IVW] OR = 3.44 per standard deviation [SD] increase, 95% confidence interval [CI] 1.66 to 7.13,  $p < 0.001$ ) and loneliness/isolation (OR = 2.98, 95% CI 1.37 to 6.46,  $p = 0.006$ ) on schizophrenia, our hypothesis that loneliness/isolation would mediate the relationship between childhood maltreatment and schizophrenia was not supported (Genomic SEM indirect effect = -0.05,  $SE = 0.05$ ,  $p = 0.255$ ; MR indirect effect = 0.10,  $SE = 0.11$ ,  $p = 0.369$ ). Furthermore, reverse mediation analysis indicated that the effect may be in the opposite direction (Genomic SEM indirect effect = 0.11,  $SE = 0.02$ ,  $p < 0.001$ ; MR indirect effect = 0.01,  $SE = 0.00$ ,  $p < 0.001$ ), accounting for 20.3%-28.9% of the total effect. The current results suggest that intervening upon loneliness/isolation in individuals with a history of childhood maltreatment is unlikely to reduce schizophrenia risk. In contrary, targeting loneliness/isolation in individuals with a genetic predisposition towards schizophrenia may diminish childhood maltreatment risk.

*Keywords:* Childhood Maltreatment, Social Isolation, Loneliness, Schizophrenia,

Mediation

**51 General Scientific Summary**

52           This study suggests that loneliness/social isolation does not mediate the relationship  
53 between childhood maltreatment and schizophrenia. Contrary to previous research, our  
54 results indicate that targeting loneliness and social isolation in individuals with a history of  
55 childhood maltreatment is unlikely to diminish schizophrenia risk. In fact, we find evidence  
56 of mediation in the reverse direction, whereby a genetic predisposition to schizophrenia  
57 increases the risk of loneliness/social isolation, which in turn increases childhood  
58 maltreatment risk.

59

60 **The Role of Loneliness and Social Isolation in Mediating the Relationship Between**  
61 **Childhood Maltreatment and Schizophrenia: A Genetically Informed Approach**

62 Childhood maltreatment, defined as the physical, sexual, or emotional abuse or  
63 neglect of a child by a care figure (Stoltenborgh, Bakermans-Kranenburg, & van IJzendoorn,  
64 2013; Stoltenborgh et al., 2011), is a major global concern affecting up to 36% of the  
65 population worldwide (Gilbert et al., 2009; Stoltenborgh et al., 2015). It is thought to be a  
66 risk factor for psychopathology in adulthood, including the development and maintenance of  
67 schizophrenia (Agnew-Blais & Danese, 2016; Bonoldi et al., 2013; Mayo et al., 2017;  
68 Velikonja et al., 2015; Warriier et al., 2021), a chronic mental health condition pathologized  
69 by persistent psychotic symptoms, including hallucinations, delusions, impaired social and  
70 interpersonal functioning (Galderisi et al., 2014; Roehr, 2013). Importantly, schizophrenia is  
71 also associated with high rates of unemployment, homelessness, and low quality of life (Eack  
72 & Newhill, 2007; Jin & Mosweu, 2017), calling for more research to better understand the  
73 risk factors for schizophrenia.

74 Observational studies have identified loneliness as a potential mediator of the  
75 relationship between childhood maltreatment and schizophrenia (Shevlin et al., 2015;  
76 Steenkamp et al., 2022). Loneliness, defined as the subjective dissatisfaction felt from one's  
77 perceived level of social connection being less than what is desired, and social isolation, the  
78 objective lack of social connections (Perlman, 1982), have prevalence rates of 10.5% (Beutel  
79 et al., 2017; operationalised as being frequently alone or having few contacts) and 20.8%  
80 (d'Hombres et al., 2021; operationalised as engaging in social activities once a month or less)  
81 respectively. Surveys indicate that individuals with schizophrenia have approximately 2.3  
82 times higher rates of loneliness than the general population (Badcock et al., 2015; Eglit et al.,  
83 2018; Stain et al., 2012). Whilst limited research has explored the role of loneliness in the  
84 association between childhood maltreatment and psychosis, it has consistently been identified

85 as a mediator of this relationship. For example, Shevlin et al. (2015) observed that loneliness  
86 mediated the relationship between childhood abuse and psychosis, even when controlling for  
87 background risk factors such as sex, age, ethnicity, cannabis use, and education. Similarly,  
88 Steenkamp et al. (2022) found that loneliness mediated the association between childhood  
89 abuse and the severity of psychosis. Comparable findings were also observed when  
90 examining social functioning and loneliness in individuals with sub-clinical psychosis levels  
91 (Boyda & McFeeters, 2015). A plausible explanation for this relationship has been proposed,  
92 whereby victims of childhood maltreatment may become sensitised to feelings of loneliness,  
93 which, as a result, could increase positive symptoms of psychosis (Steenkamp et al., 2022).

94         If loneliness does mediate the association between maltreatment and schizophrenia,  
95 targeted interventions aimed at reducing feelings of loneliness/isolation could be offered to  
96 victims of childhood maltreatment, as a preventative measure for later life psychosis.  
97 However, limitations with observational studies (e.g., unmeasured confounding and reverse  
98 causation) have meant that the robustness of this mediatory relationship has thus far not been  
99 explored. Employing genetically informed methods, such as Mendelian randomisation – that  
100 are less susceptible to confounding, reverse causation, and measurement error – may help to  
101 address these challenges and provide insights into putative causal effects (Davey Smith &  
102 Ebrahim, 2005; Pingault et al., 2018). This is because (1) genetic variants are fixed at  
103 conception and therefore cannot be modified by disease onset, and (2) genetic variants are  
104 randomly allocated during meiosis, which minimizes confounding and measurement biases  
105 when these are being used as instrumental variables for modifiable risk factors (Davey Smith  
106 et al., 2007; Pingault et al., 2018).

107         As such, to address limitations of observational research, the current study  
108 investigated the extent to which the relationship between childhood maltreatment and  
109 schizophrenia is mediated by a combined phenotype of loneliness and social isolation

110 (hereafter loneliness/isolation) using two complementary, genetically informed methods:  
111 genomic structural equation modelling and multivariable Mendelian randomisation. We  
112 hypothesised that the relationship between childhood maltreatment and schizophrenia would  
113 be mediated by loneliness/isolation.

## 114 **Methods**

### 115 **Sources of Data**

116 Ethical approval was granted by the local ethics committee (Project: 21-108). Openly  
117 accessible genome-wide summary statistics were selected for each trait from the most recent  
118 and largest GWASs at the time of data analysis (see Table 1). To reduce the likelihood of  
119 false positive results due to differing allele frequencies across admixed populations (Caliebe  
120 et al., 2022; Patterson et al., 2006), only studies of European ancestry were selected.

121 **Childhood Maltreatment** summary statistics were obtained from Warrier et al.'s  
122 (Warrier et al., 2021) meta-analysis, which included GWAS summary statistics from the  
123 Psychiatric Genomics Consortium (PGC;  $n = 26,290$ ) as well as individual-level data from  
124 four cohort studies: the UK Biobank (UKBB;  $n = 143,473$ ), Avon Longitudinal Study of  
125 Parents and Children (ALSPAC;  $n = 8346$ ), Adolescent Brain Cognitive Development Study  
126 (ABCD;  $n = 5400$ ), and Generation R (GenR;  $n = 1905$ ). Childhood maltreatment was  
127 defined as emotional and physical neglect, and emotional, physical, and sexual abuse.  
128 Prospective reports of childhood maltreatment were largely reported by a parent or caregiver,  
129 whereas retrospective reports were self-reported. Mean age ranged from 10-64 years.  
130 Association testing was conducted using linear mixed-effects models in the UKBB,  
131 ALSPAC, and ABCD, which accounted for both relatedness and population stratification.  
132 Additionally, in the UKBB and ABCD, 20 genetic principal components (PCs) were included  
133 to further control for population stratification. In GenR, a linear regression was employed for  
134 association analyses, with five PCs included as covariates.

135           **Loneliness/Isolation** summary statistics were obtained from Day, Ong, and Perry's  
136 (2018) multi-trait GWAS (MTAG), which looked at a combined loneliness and social  
137 isolation phenotype using UK Biobank data (effective  $N = 487,647$ ). Loneliness/isolation  
138 was defined based on self-reported responses to the following questions: (1) 'Do you often  
139 feel lonely?'; (2) 'Including yourself, how many people are living together in your  
140 household?' and 'How often do you visit friends or family or have them visit you?'; and (3)  
141 'How often are you able to confide in someone close to you?'. Cases were defined as those  
142 who (a) perceived themselves as lonely, (b) lived alone and never visited family or friends  
143 outside their household, or (c) could never or almost never confide in someone. The multi-  
144 trait GWAS combined summary statistics from three GWASs based on each of these three  
145 questions. Relatedness and cryptic population structure were accounted for by using a linear  
146 mixed model in association analyses.

147           **Schizophrenia** summary statistics for autosomal SNPs were obtained from  
148 Trubetskoy et al. (2022) PGC Wave 3 GWAS meta-analysis of European ancestry cohorts  
149 (53,386 cases and 77,258 controls). Cases were defined as individuals with schizophrenia,  
150 schizoaffective disorder, or schizophrenia spectrum disorder. Controls consisted of screened  
151 and unscreened individuals for schizophrenia and other psychoses. Cases were recruited from  
152 community-based samples, hospital settings with ascertainment for clozapine treatment, or  
153 mixed (i.e., community/hospital) setting. Diagnostic strategy included research diagnostic  
154 interview, review of medical records, consensus diagnosis, and mixed strategy. Detailed  
155 diagnostic strategy and case-control sample description for approximately 75 European  
156 cohorts included in this meta-analysis are available in the original study (Trubetskoy et al.,  
157 2022). Within each cohort, a subset of the first 20 genetic PCs were included as covariates in  
158 GWAS analyses. By default, all cohorts included the first four PCs and any additional PC  
159 associated with case-control status.



**160 Genetic Mediation using Genomic SEM**

161 To perform a genetic mediation analysis, we used genomic structural equation  
162 modelling (Genomic SEM; Grotzinger et al., 2019), which uses summary statistics to model  
163 the genetic covariance structure of complex traits. The data were analysed in R (v4.0.3; R  
164 Core Team, 2020) as per developer instructions  
165 (<https://gist.github.com/MichelNivard/04bf4ddcf3c32f905175de3058ca967a>).

166 First, summary statistics underwent quality control, whereby only high-quality,  
167 common single nucleotide polymorphisms (SNPs) (INFO > 0.9 with minor allele frequency >  
168 0.01) that were also available in HapMap3 reference panel were retained. The formatted files  
169 were used as input to Linkage Disequilibrium (LD) score regression (Bulik-Sullivan et al.,  
170 2015a,b), which estimated the genetic correlation between each pair of traits and returned the  
171 genetic covariance and its associated sampling covariance matrices. Subsequently, the genetic  
172 mediation model was specified using diagonally weighted least squares estimation  
173 (Grotzinger et al., 2019), whereby childhood maltreatment was included as the predictor,  
174 schizophrenia as the outcome and loneliness/isolation as the mediator.

175 Mediation analysis enabled us to decompose the total effect of childhood  
176 maltreatment on schizophrenia into direct and indirect effects. The direct effect represented  
177 the effect of childhood maltreatment on schizophrenia controlling for loneliness/isolation,  
178 whereas the indirect effect captured the portion of the total effect that is mediated by  
179 loneliness/isolation. To account for the presence of potential bidirectional effects, we ran a  
180 subsequent genetic mediation, in which schizophrenia was included as the predictor,  
181 loneliness/isolation as the mediator, and childhood maltreatment as the outcome.

**182 Genetic Mediation using Mendelian Randomisation**

183 To gain insights into putatively causal effects, we performed Mendelian  
184 randomisation (MR) mediation analysis using the TwoSample MR package (version 0.5.6;

185 Hemani, Tilling, & Davey Smith, 2017). MR is a causal inference technique that uses genetic  
186 variants as instrumental variables for an exposure of interest. To ensure genetic variants are  
187 valid instrumental variables, three key MR assumptions need to be satisfied: (1) genetic  
188 variants should be associated with the exposure, (2) genetic variants should not be associated  
189 with any confounders of the exposure-outcome relationship, (3) genetic variants should not  
190 affect the outcome via any other path than the modelled exposure (i.e., there should be no  
191 horizontal pleiotropy) (Lawlor et al., 2008). The latter assumption is what sets MR analysis  
192 apart from Genomic SEM. Specifically, unlike Genomic SEM which uses genetic covariance  
193 structure to analyse pleiotropy involving multiple genes and traits, MR assumes the absence  
194 of horizontal pleiotropy and includes sensitivity analyses, such as MR-Egger intercept test, to  
195 test for it. The advantage of using MR over conventional observational epidemiology  
196 techniques is that it offers some protection from reverse causality and unmeasured  
197 confounding (Davey Smith, 2004). In the present study, we used two-sample univariable MR  
198 to estimate the *total effect* of childhood maltreatment on (i) schizophrenia and (ii)  
199 loneliness/isolation (i.e., path *a*; Figure 1a). Two-sample multivariable MR was used to  
200 estimate the *direct effect* of childhood maltreatment on schizophrenia controlling for  
201 loneliness/isolation, and the *direct effect* of loneliness/isolation on schizophrenia controlling  
202 for childhood maltreatment (i.e., path *b*; Figure 1b). The total effect of loneliness/isolation on  
203 schizophrenia was also estimated, although it was not of primary interest. All analyses were  
204 performed in R (version 4.2.1; R Core Team, 2020). Post hoc power calculation for the MR  
205 analyses was conducted separately for each exposure-outcome combination using a web-  
206 based tool (<https://sb452.shinyapps.io/power/>; Burgess, 2014). Power to detect an indirect  
207 effect was obtained by multiplying the power for path *a* (maltreatment->loneliness/isolation)  
208 and path *b* (loneliness/isolation->schizophrenia) (Kenny & Judd, 2014).

**Commented [VB1]:** added path *a* (unchanged analysis, previously reported in the supplements)

**Commented [VB2]:** unchanged analysis, added for clarification

209 **Instrument Selection.** In univariable MR, we selected 14 independent ( $r^2 < 0.1$   
210 within a 1000 kb window size, as defined by the authors), genome-wide significant SNPs ( $p$   
211  $< 5 \times 10^{-8}$ ) associated with childhood maltreatment from the GWAS by Warrier et al. (2021)  
212 and extracted these same SNPs from the schizophrenia outcome GWAS (Trubetskoy et al.,  
213 2022). As four of the 14 SNPs were not available in the outcome GWAS and one of the 14  
214 SNPs was palindromic with intermediate allele frequencies, our final set of instruments  
215 included nine SNPs. For loneliness/isolation, all 15 independent SNPs ( $p < 5 \times 10^{-8}$ ) that were  
216 associated with loneliness/isolation (Day et al., 2018) were extracted from the outcome  
217 schizophrenia GWAS. In multivariable MR, we combined genome-wide significant SNPs  
218 associated with childhood maltreatment or loneliness/isolation into a single set of  
219 instruments, resulting in 24 SNPs. As only independent variants were retained (using default  
220 multivariable MR settings of  $r^2 < 0.001$  within a 10,000 kb window size), one of the 24 SNPs  
221 was removed due to being in LD with another variant. This final set of 23 SNPs was then  
222 extracted from the schizophrenia outcome GWAS (Trubetskoy et al., 2022). SNP effects for  
223 the exposure and outcome data were harmonised to correspond to the same effect allele.  
224 When dealing with palindromic SNPs, we attempted to infer positive strand alleles. When  
225 this was not possible, SNPs with intermediate allele frequencies ( $MAF > 0.42$ ) were  
226 excluded.

227 **Univariable MR.** To estimate the genetically predicted effect of childhood  
228 maltreatment on (1) schizophrenia and (2) loneliness/isolation, and (3) the genetically  
229 predicted effect of loneliness/isolation on schizophrenia, we carried out a two-sample  
230 univariable MR analysis using the inverse variance weighing (IVW) method (Burgess,  
231 Butterworth, & Thompson, 2013) separately for each exposure-outcome combination. These  
232 were followed by a range of sensitivity analyses, including MR Egger (Bowden, Davey  
233 Smith, & Burgess, 2015), the weighted median (Bowden, Davey Smith, Haycock, & Burgess,

Commented [VB3]: adding path  $\alpha$

234 2016), the simple mode (Hartwig et al., 2017), the weighted mode (Burgess, Foley, Allara,  
235 Staley, & Howson, 2020) and the leave-one-out (Burgess, Bowden, Fall, Ingelsson, &  
236 Thompson, 2017) methods. The resulting logistic betas for schizophrenia risk were  
237 exponentiated to obtain odds ratios (OR). Sensitivity analyses assessed the robustness of our  
238 findings to the potential bias of horizontal pleiotropy, where the genetic variants affect the  
239 risk of schizophrenia through other pathways than the modelled exposure. A discrepancy in  
240 effect estimates from the primary IVW analysis and the range of sensitivity analyses were  
241 indicative of potential violations of the MR assumptions. Instrument strength was assessed  
242 using the  $F$ -statistic, where  $F < 10$  was indicative of weak instruments (Burgess et al., 2011).  
243 Heterogeneity of genetic variants was evaluated using Cochran's  $Q$  statistic (Bowden et al.,  
244 2019).

245 **Multivariable MR.** Next, we used multivariable MR to estimate (1) the direct effect  
246 of childhood maltreatment on the risk of schizophrenia while controlling for  
247 loneliness/isolation and (2) the direct effect of loneliness/isolation on schizophrenia,  
248 controlling for childhood maltreatment. In multivariable MR, genetic variants can be  
249 associated with more than one exposure, which enables the estimation of direct effects of  
250 each exposure, conditional on the other modelled exposures (Burgess & Thompson, 2015;  
251 Carter et al., 2021). As before, MR-Egger regression was used as a sensitivity analysis  
252 (Bowden et al., 2015) using the Mendelian randomisation package (version 0.6.0; Yavorska  
253 & Burgess, 2017). Instrument strength was evaluated using a conditional  $F$ -statistic and  
254 heterogeneity was assessed via a modified form of the Cochran's  $Q$  statistic (Sanderson et al.,  
255 2021), estimated using the MVMR package (version 0.3) with default settings (Sanderson,  
256 Davey Smith, Windmeijer, & Bowden, 2019). Logistic betas for schizophrenia risk were  
257 exponentiated.

Commented [VB4]: unchanged analysis, added for clarification

258 **Mediation Analysis.** To estimate the degree to which the effect of childhood  
259 maltreatment on schizophrenia risk is mediated by loneliness/isolation (i.e., the indirect  
260 effect), we applied the product of coefficients method (Burgess et al., 2015) by multiplying  
261 the total effect of childhood maltreatment on loneliness/isolation obtained in univariable MR  
262 (i.e., path *a*) with the direct effect of loneliness/isolation on schizophrenia obtained in  
263 multivariable MR (i.e., path *b*). The standard errors (*SE*) for the indirect effect were estimated  
264 using the Delta method (Thompson et al., 2016); see Methods section of the Supplementary  
265 Material). Note that the product of coefficients method was applied on the non-exponentiated  
266 coefficients.

267 **Reverse MR.** We tested for the presence of bidirectional effects by repeating the  
268 above analyses in the reverse direction. This permitted us to estimate the genetically  
269 predicted effect of schizophrenia on childhood maltreatment and loneliness/isolation, as well  
270 as to examine the reverse mediation pathway. Instrument selection for reverse MR is detailed  
271 in the Methods section of the Supplementary Material.

### 272 **Transparency and Openness**

273 All data and code on which study conclusions are based are publicly available and  
274 may be accessed via links included in the Data and Code Availability section. The data were  
275 analysed in R (version 4.0.3; R Core Team, 2020) using Genomic SEM (version 0.0.5;  
276 Grotzinger et al., 2019), TwoSampleMR (version 0.5.6; Hemani, Tilling, & Davey Smith,  
277 2017) and MVMR (version 0.3; Sanderson, Davey Smith, Windmeijer, & Bowden, 2019)  
278 packages. This study was not preregistered. Where applicable, JARS guidelines (Kazak,  
279 2018) were followed.

### 280 **Results**

281 All three phenotypes were genetically correlated. Childhood maltreatment was  
282 strongly correlated with loneliness/isolation ( $r_g = 0.64$ ,  $SE = 0.04$ ,  $p < 0.001$ ) and moderately

Commented [VB5]: Updated to product of coefficients method

283 correlated with schizophrenia ( $r_g = 0.38$ ,  $SE = 0.03$ ,  $p < 0.001$ ). In contrast, only a weak  
284 genetic correlation was observed between loneliness/isolation and schizophrenia ( $r_g = 0.19$ ,  
285  $SE = .03$ ,  $p < 0.001$ ).

#### 286 **Genetic Mediation using Genomic SEM**

287 Genetic mediation analysis revealed a significant total effect of childhood  
288 maltreatment on schizophrenia (total effect = 0.38,  $SE = 0.04$ ,  $p < 0.001$ ; Figure 2).  
289 Childhood maltreatment was also positively associated with loneliness/isolation ( $a$  path =  
290 0.43,  $SE = 0.05$ ,  $p < 0.001$ ), while loneliness/isolation showed no association with  
291 schizophrenia when controlling for childhood maltreatment ( $b$  path = -0.08,  $SE = 0.05$ ,  $p =$   
292 0.108). The mediating effect of loneliness/isolation was not significant (indirect effect = -  
293 0.05,  $SE = 0.05$ ,  $p = 0.255$ ). In line with this non-significant indirect effect, we continued to  
294 observe a positive association between childhood maltreatment and schizophrenia after  
295 controlling for loneliness/isolation (direct effect = 0.43,  $SE = 0.07$ ,  $p < 0.001$ ). Taken  
296 together, these results suggest that loneliness/isolation does not mediate the relationship  
297 between childhood maltreatment and schizophrenia (Figure 2).

298 In contrast to the findings above, mediation analysis in the reverse direction identified  
299 a significant effect of genetic liability to schizophrenia on childhood maltreatment (total  
300 effect = 0.38,  $SE = 0.03$ ,  $p < 0.001$ ) that attenuated after adjusting for loneliness/isolation  
301 (direct effect = 0.27,  $SE = 0.04$ ,  $p < 0.001$ ). Additionally, the mediating effect of  
302 schizophrenia liability on childhood maltreatment via loneliness/isolation was significant  
303 (indirect effect = 0.11,  $SE = 0.02$ ,  $p < 0.001$ ), with 28.9% of the total effect being mediated  
304 by loneliness/isolation (Supplementary Figure 1).

#### 305 **Genetic Mediation using Mendelian Randomisation**

306 Results from univariable MR revealed an effect of childhood maltreatment on  
307 schizophrenia ( $OR_{IVW} = 3.44$  per standard deviation [SD] increase, 95% confidence interval

308 [CI] 1.66 to 7.13,  $p < 0.001$ ), with consistent direction of effect estimates across all  
309 sensitivity analyses (Figure 3; Supplementary Table 1; Supplementary Figure 2). While MR-  
310 Egger showed no evidence of horizontal pleiotropy (intercept = -0.09,  $SE = 0.08$ ,  $p = 0.331$ ),  
311 some heterogeneity was detected ( $Q_{IVW} = 56.51$ ,  $p < 0.001$ ). A significant effect of  
312 loneliness/isolation on schizophrenia was also observed ( $OR_{IVW} = 2.98$  per SD increase, 95%  
313 CI 1.37 to 6.46,  $p = 0.006$ ), with consistent effect estimates across all sensitivity analyses  
314 apart from MR-Egger, which was not significant and in the opposite direction (Figure 3;  
315 Supplementary Table 1, Supplementary Figure 3). As before, MR-Egger showed no evidence  
316 of horizontal pleiotropy (intercept = 0.05,  $SE = 0.03$ ,  $p = 0.103$ ). However, there was  
317 heterogeneity among the genetic variants ( $Q_{IVW} = 69.04$ ,  $p < 0.001$ ), suggesting that one or  
318 more of the loneliness/isolation-associated variants may be pleiotropic. The effect of  
319 childhood maltreatment on loneliness/isolation (i.e., path  $a$ ) was also significant ( $b_{IVW} = 0.23$   
320 per SD increase, 95% CI 0.13 to 0.32,  $p < 0.001$ ), albeit only partly supported by the  
321 sensitivity analyses (see Results section of the Supplementary Material; Supplementary Table  
322 1; Supplementary Figure 4). Leave-one-out analysis suggested that no single SNP was  
323 driving the IVW estimates for the above analyses (Supplementary Figures 5-7). No evidence  
324 of weak instrument bias was detected (maltreatment  $F = 40.39$ ; loneliness = 38.75;  
325 Supplementary Tables 2 and 3).

326 Results from multivariable MR revealed a direct effect of childhood maltreatment on  
327 schizophrenia ( $OR_{IVW} = 3.49$  per SD increase, 95% CI 1.75 to 6.95,  $p < 0.001$ ), which was  
328 comparable in magnitude to its total effect ( $OR_{IVW} = 3.44$ ). The MR-Egger estimate for the  
329 direct effect was consistent with the estimate from the IVW analysis ( $OR = 3.98$  per SD  
330 increase, 95% CI 0.73 to 21.58,  $p = 0.109$ ), albeit not significant. No evidence of horizontal  
331 pleiotropy was detected as the MR-Egger intercept was close to zero (intercept = -0.00,  $SE =$   
332 0.03,  $p = 0.867$ ). Conditional  $F$ -statistics for childhood maltreatment and loneliness/isolation

**Commented [VB6]:** unchanged results (Table S7), just reporting path  $a$  here for clarification (hence subsequent figure numbers changed)

333 were 10.16 and 12.27 respectively, suggesting no bias due to weak instruments. However, the  
334 modified Cochran's Q statistic indicated presence of heterogeneity among the genetic  
335 variants ( $Q_{A-IVW} = 86.21, p < 0.001$ ).

336 Compared to the total effect of loneliness/isolation on schizophrenia, the direct effect  
337 of loneliness/isolation (i.e., *b* path) attenuated from an OR of 2.98 to 1.55 (95% CI 0.61 to  
338 3.94,  $p = 0.360$ ), suggesting that the effect of loneliness/isolation was not independent from  
339 childhood maltreatment. Correspondingly, the multivariable MR-Egger estimate was also not  
340 significant (OR = 1.95 per SD increase, 95% CI 0.11 to 35.34,  $p = 0.650$ ) (Figure 3;  
341 Supplementary Table 1).

342 Mediation analysis using the product of coefficients method revealed that, in line with  
343 findings using Genomic SEM, the indirect effect of childhood maltreatment was not  
344 significant (indirect effect = 0.10,  $SE = 0.11, p = 0.369$ ), indicating a lack of mediation.

345 **Reverse MR.** Results from univariable MR in the reverse direction revealed a modest  
346 effect of schizophrenia liability on childhood maltreatment ( $b_{IVW} = 0.05$  per SD increase,  
347 95% CI 0.03 to 0.06,  $p < 0.001$ ), with relatively consistent evidence across sensitivity  
348 analyses (Figure 4; Supplementary Table 4; Supplementary Figure 8). The IVW estimate for  
349 the effect of schizophrenia on loneliness/isolation (path *a*) was also significant ( $b_{IVW} = 0.02$   
350 per SD increase, 95% CI 0.01 to 0.03,  $p < 0.001$ ); however, it was not robustly supported by  
351 the sensitivity analyses (Figure 4; Supplementary Table 4; Supplementary Figure 9). Results  
352 for the remaining sensitivity analyses are detailed in the Results section of the Supplementary  
353 Material.

354 In MVMR analysis, we observed a direct effect of schizophrenia liability on  
355 childhood maltreatment after controlling for loneliness/isolation ( $b_{IVW} = 0.04$  per SD  
356 increase, 95% CI 0.02 to 0.05,  $p < 0.001$ ; Figure 4), which was slightly smaller than the total  
357 effect ( $OR_{IVW} = 0.05$ ); and a direct effect of loneliness/isolation on childhood maltreatment

Commented [VB7]: Updated to product of coefficients method



358 after controlling for schizophrenia liability (b<sub>rvw</sub> = 0.44 per SD increase, 95% CI 0.26 to  
359 0.62,  $p < 0.001$ ; Supplementary Table 4). Sensitivity analyses for MVMR are reported in the  
360 Results section of the Supplementary Material. In line with findings from Genomic SEM,  
361 mediation analysis revealed a significant mediating effect of schizophrenia on childhood  
362 maltreatment via loneliness/isolation (indirect effect = 0.01,  $SE = 0.00$ ,  $p < 0.001$ ), which  
363 accounted for 20.3% of the total effect. Note that alternative mediation analysis  
364 (schizophrenia → childhood maltreatment → loneliness/isolation) was not performed given  
365 the lack of robust support for a total effect of schizophrenia on loneliness/isolation.

366 **Post hoc power calculation.** At an alpha level of 0.05, we had over 80% power to  
367 detect small-to-medium sized effects in our univariable MR analyses. Our power to detect an  
368 indirect effect from childhood maltreatment to schizophrenia via loneliness/isolation was  
369 93.4% (calculated given 100% power for path *a* and 93.4% power for path *b*). Power for each  
370 exposure-outcome combination and detailed information regarding input parameters is  
371 provided in Supplementary Table 7.

### 372 Discussion

373 The current study explored the degree to which loneliness/isolation mediates the  
374 relationship between childhood maltreatment and schizophrenia using two genetically-  
375 informed methods: Genomic SEM and Mendelian randomisation. Positive genetic  
376 correlations were observed between all three traits, suggesting a shared genetic underpinning.  
377 However, contrary to our hypothesis, we found no support for the mediating role of  
378 loneliness/isolation, indicating that loneliness/isolation is not a robust mediatory factor. These  
379 findings suggest that the development of schizophrenia, which might be exacerbated by early  
380 life abuse or neglect, is unlikely to be explained through loneliness/isolation. In contrast, we  
381 find evidence of mediation in the reverse direction, whereby a genetic predisposition to

**Commented [VB8]:** unchanged results, just added for clarification

**Commented [VB9]:** Updated to product of coefficients method (changed from 19.5% to 20.3%)

382 schizophrenia might increase loneliness/isolation, which in turn could increase childhood  
383 maltreatment risk.

384         In line with a growing body of literature (Brown et al., 2014; Croft et al., 2019;  
385 Schaefer et al., 2018; South et al., 2015), the present study found a bidirectional direct effect  
386 between childhood maltreatment and schizophrenia, indicating that experiencing early life  
387 abuse or neglect may increase risk of developing schizophrenia in later life. This also  
388 corroborated findings from a recent Mendelian randomisation study that reported a  
389 bidirectional effect between childhood maltreatment and schizophrenia (Warrier et al., 2021).  
390 Additionally, a strong genetic association between childhood maltreatment and  
391 loneliness/isolation was observed, which aligned with existing observational research, in  
392 which victims of childhood maltreatment were found to present with lower levels of  
393 interpersonal functioning due to mistrust and suspicion of others (Blanchard-Dallaire &  
394 Hébert, 2014; Boyda & McFeeters, 2015). It is plausible that early life trauma predisposes  
395 individuals to develop negative schemas about others, which can increase feelings of mistrust  
396 (Cukor & McGinn, 2006; Karatzias et al., 2016; Tezel et al., 2015), and may in turn  
397 compromise the development of long-lasting relationships (Blanchard-Dallaire & Hébert,  
398 2014). Clinical interventions should therefore aim to reduce feelings of loneliness and social  
399 isolation for individuals who have experienced childhood maltreatment (e.g., through  
400 community groups).

401         Unexpectedly, when controlling for childhood maltreatment, we observed no  
402 relationship between loneliness/isolation and schizophrenia. This finding contradicts previous  
403 research, which reports a positive – and potentially causal – association between the two  
404 traits (Andreu-Bernabeu et al., 2022; Boyda & McFeeters, 2015; Steenkamp et al., 2022).  
405 The discrepancy in findings may be driven by differences in the type of data analysed (i.e.,  
406 genetic instead of phenotypic), the type of mediator used (e.g., loneliness/isolation as

407 opposed to difficulties in social engagement), and/or the choice of covariates (i.e., controlling  
408 for childhood maltreatment or not). Our multivariable MR results indeed indicate that  
409 covarying factors might play an important role. Specifically, in line with existing research  
410 (Andreu-Bernabeu et al., 2022), univariable MR revealed a unidirectional effect of  
411 loneliness/isolation on schizophrenia. However, this effect largely disappeared when  
412 controlling for childhood maltreatment, suggesting that the relationship between  
413 loneliness/isolation and schizophrenia is not independent from childhood maltreatment.

414       Most importantly, the hypothesis that loneliness/isolation would mediate the  
415 relationship between childhood maltreatment and schizophrenia was not supported when  
416 using Genomic SEM or MR. This finding contradicted previous phenotypic research which  
417 argues that loneliness/isolation can mediate the relationship between childhood maltreatment  
418 and schizophrenia (Boyda & McFeeters, 2015; Shevlin et al., 2015; Steenkamp et al., 2022),  
419 and suggests instead that the emergence of schizophrenia, which may be aggravated by early  
420 life maltreatment, is unlikely to be explained through loneliness/isolation. While previous  
421 observational studies found evidence of mediation, they focused on different aspects of  
422 psychosis and used a variety of different measures and definitions of childhood maltreatment.  
423 For instance, two studies only considered positive symptoms of psychosis rather than  
424 including the whole spectrum of symptoms commonly observed in schizophrenia (Boyda &  
425 McFeeters, 2015; Steenkamp et al., 2022), while another only assessed the effects of physical  
426 and sexual abuse (Shevlin et al., 2015), instead of including a more comprehensive measure  
427 of childhood maltreatment that also accounts for emotional abuse and/or physical and  
428 emotional neglect. Therefore, it is difficult to conclude to what extent differences in findings  
429 were driven by the heterogeneity in study design (e.g., genetically informed vs phenotypic  
430 designs) and/or differences in the type of measures used.

431 Interestingly, mediation analysis in the reverse direction identified loneliness/isolation  
432 as lying on the pathway from schizophrenia liability to childhood maltreatment, with up to  
433 29% of this relationship explained by loneliness/isolation. One way to interpret this finding is  
434 that individuals with a genetic predisposition toward schizophrenia may experience  
435 difficulties in forming and maintaining social connections. This may in turn lead to social  
436 isolation, fewer protective factors (e.g., making individuals less likely or less able to seek  
437 help in at-risk or abusive environments – and less likely to be identified by others as someone  
438 in need of help) and consequently a greater risk for childhood maltreatment. Gene–  
439 environment correlations provide alternative explanations: parents with a genetic  
440 predisposition toward schizophrenia may create an environment for their children that  
441 increases the risk of loneliness/isolation and childhood maltreatment (i.e., passive gene–  
442 environment correlation); children’s genetic make-up influences their behaviours and choice  
443 of environments, which may foster loneliness/isolation and/or childhood maltreatment (i.e.,  
444 active gene–environment correlation); and, children’s genetic make-up impacts their  
445 behaviour and characteristics, evoking reactions from others to that behaviour or  
446 characteristic (i.e., reactive gene–environment correlation) (Griffiths, 2005; Jaffee & Price,  
447 2007; Knafo & Jaffee, 2013; Plomin et al., 1977; Warrier et al., 2021).

448 Our study has to be seen in light of the following limitations. First, due to  
449 methodological constraints, the current findings cannot be extrapolated to non-European  
450 ancestry individuals. Secondly, the present study only modelled the shared genetic  
451 architecture between childhood maltreatment, loneliness/isolation, and schizophrenia.  
452 Additional genetic and phenotypic confounds (e.g., education, cannabis use, adult  
453 victimisation) were not considered, which could inform our understanding of the gene-  
454 environment correlations that may increase schizophrenia risk (Hernán et al., 2004; Lutz et  
455 al., 2017; Warrier et al., 2021). Third, as alleles are assigned randomly only within but not

456 across families, the important MR assumption of no genetic confounding may be violated.  
457 However, as all GWASs included here controlled for population stratification, we consider  
458 this assumption to likely be upheld.

459       Lastly, our ability to draw strong causal inferences from our results is limited. This  
460 largely stems from the difficulty in disentangling causality from pleiotropic effects, as  
461 observed associations may arise due to multiple scenarios, namely: (i) independent genetic  
462 effects on trait 1 and 2 (i.e., horizontal pleiotropy), (ii) genetic effects on trait 1, which then  
463 causes trait 2 (the basis of MR), (iii) and/or genetic effects on trait 2, which then causes trait 1  
464 (i.e., reverse MR). Considering that pleiotropy was not detected by our sensitivity analyses, it  
465 is possible that our findings are driven by either one of the latter two scenarios. Nonetheless,  
466 drawing strong causal environmental inferences from MR would be inappropriate, as the  
467 crucial assumption on the absence of horizontal pleiotropy cannot be completely ruled out.

468       Despite these limitations, the fact that our results were consistent across two divergent  
469 methods (Genomic SEM and MR) provides greater confidence in the robustness of our  
470 findings. Future studies should investigate alternative mediatory personality (e.g.,  
471 neuroticism) and lifestyle factors (e.g., smoking) in the relationship between childhood  
472 maltreatment and schizophrenia to inform intervention-based treatment programmes. While  
473 observational studies have already identified factors such as perceived discrimination,  
474 emotion dysregulation, mood symptoms, and the lack of social support as potential mediators  
475 in the association between childhood abuse and psychosis (Sideli et al., 2020), little has been  
476 done to assess the robustness of these relationships using genetically informed techniques.

477       To the best of our knowledge, the current study is the first to explore the mediating  
478 role of loneliness/isolation in the relationship between childhood maltreatment and  
479 schizophrenia using genetic data. Using high powered GWAS summary statistics ( $N =$   
480 105,318 - 487,647), our results suggest that loneliness/isolation does not mediate the

481 association between childhood maltreatment and schizophrenia. In fact, we find evidence of  
482 mediation in the opposite direction. As such, whilst there may be merit in clinical  
483 interventions targeting loneliness/isolation in individuals who have experienced childhood  
484 maltreatment (e.g., social conditioning interventions; Shevlin et al., 2015), the current results  
485 suggest that targeting loneliness/isolation in this group is unlikely to reduce the risk of  
486 developing schizophrenia. Greater benefits may be observed by intervening upon  
487 loneliness/isolation in individuals with a genetic predisposition towards schizophrenia, as this  
488 may diminish childhood maltreatment risk. Taken together, these findings provide key novel  
489 insights into the link between childhood maltreatment and schizophrenia, with significant  
490 implications for targeted clinical interventions.

491

#### 492 **Data and Code Availability**

493 All datasets analysed in the present study are publicly available. Summary statistics  
494 for childhood maltreatment and loneliness/social isolation can be obtained from the  
495 University of Cambridge Apollo repository (childhood maltreatment:  
496 <https://www.repository.cam.ac.uk/handle/1810/318326>; loneliness/social isolation:  
497 <https://www.repository.cam.ac.uk/handle/1810/277812>). Summary statistics for  
498 schizophrenia can be obtained from the Walters group data repository  
499 (<https://walters.psych.cf.ac.uk>). Analysis code for this study has been made publicly  
500 available on GitHub and can be accessed at <https://github.com/VilteBaltra/loneliness->  
501 [mediation](https://github.com/VilteBaltra/loneliness-meditation).

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**Table 1***Characteristics of the Contributing Genome-wide Association Studies*

<b>Phenotype</b>	<b>Assessment</b>	<b>N</b>	<b>Age</b>	<b>GWAS</b>	<b><math>h^2_{\text{SNP}}</math></b>
<b>Childhood Maltreatment</b>	Self-report / Parent-report	185,414	10-64	Warrier et al. (2021)	8%
<b>Loneliness/Isolation</b>	Self-report	487,647	40-69	Day et al. (2018)	4.2%
<b>Schizophrenia</b>	*Multiple sources	130,644	NA	Trubetskoy et al. (2022)	24%

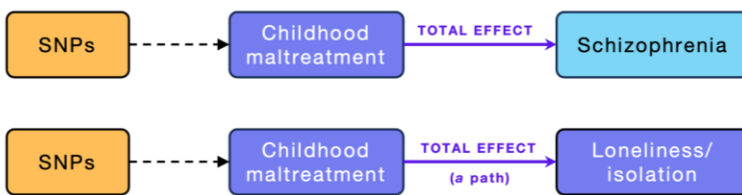
*Note.*  $h^2_{\text{SNP}}$  = SNP heritability reported in the respective GWASs; *N* = sample size; Age = mean age in years across contributing cohorts; GWAS = genome-wide association study.

\*Cases were ascertained based on multiple sources of information such as medical records, semi-structured interviews, and clinical evaluation by psychiatrists. More detailed information on assessment methods is included in the original study by Trubetskoy et al. (2022).

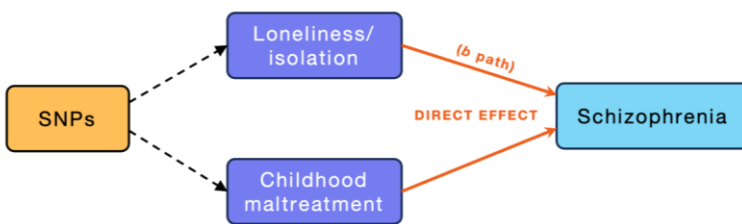
**Figure 1**

*Estimation of the Total and Direct Effects from Mendelian Randomisation Analysis*

a) Univariable MR



b) Multivariable MR

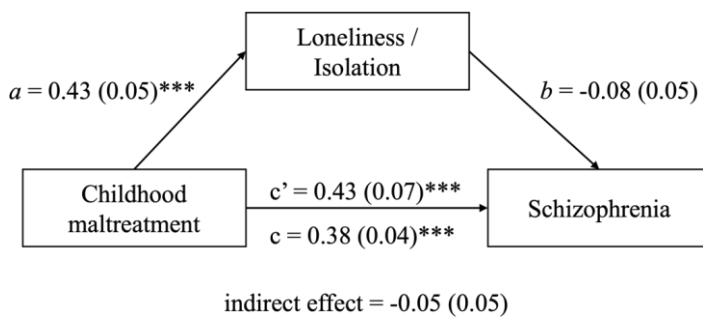


*Note.* The purple arrow represents the total effect of childhood maltreatment (exposure) on schizophrenia (outcome) and the total effect on loneliness/isolation (mediator; path *a*).

Orange arrows represent (i) the direct effect of childhood maltreatment on schizophrenia and (ii) the direct effect of loneliness/isolation on schizophrenia (path *b*). MR = Mendelian randomisation; SNPs = single nucleotide polymorphisms.

**Figure 2**

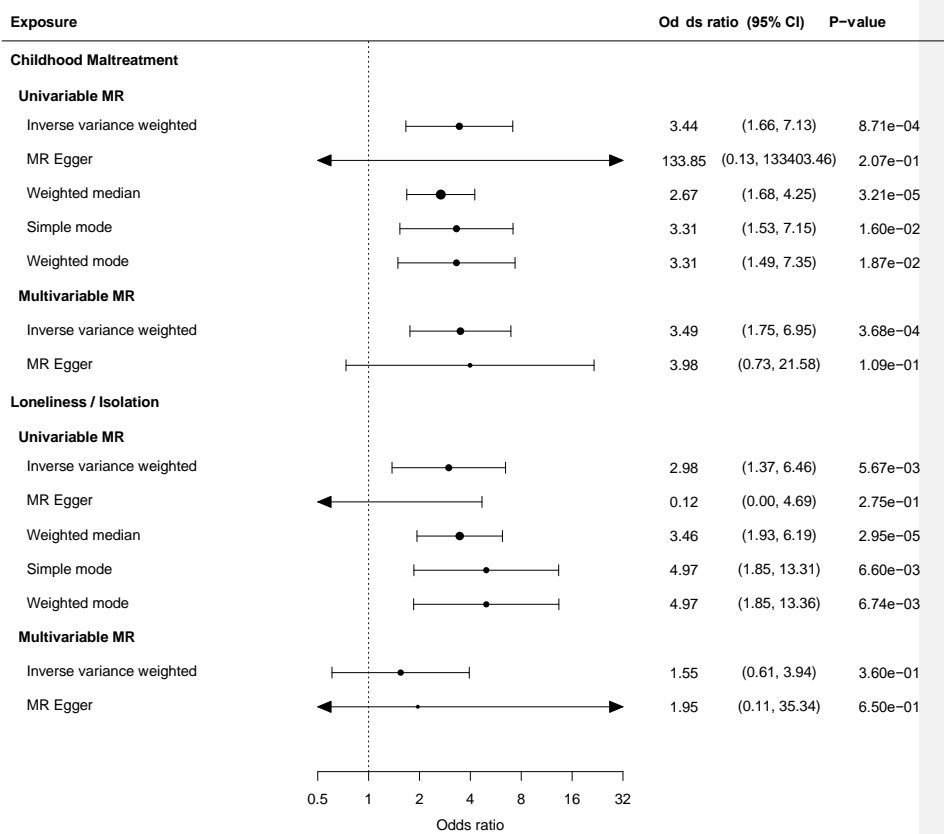
*Genetic Mediation Analysis Estimating the Mediating Effect of Loneliness/Isolation in the Relationship between Childhood Maltreatment and Schizophrenia*



*Note.* Standardised regression weights and the corresponding standard errors (in parentheses) are reported. *a* path = effect of the predictor (childhood maltreatment) on the mediator (loneliness/isolation); *b* path = effect of the mediator (loneliness/isolation) on the outcome (schizophrenia); *c'* path = direct effect of childhood maltreatment on schizophrenia; *c* path = total effect of childhood maltreatment on schizophrenia. \*\*\* *p* < 0.001.

**Figure 3**

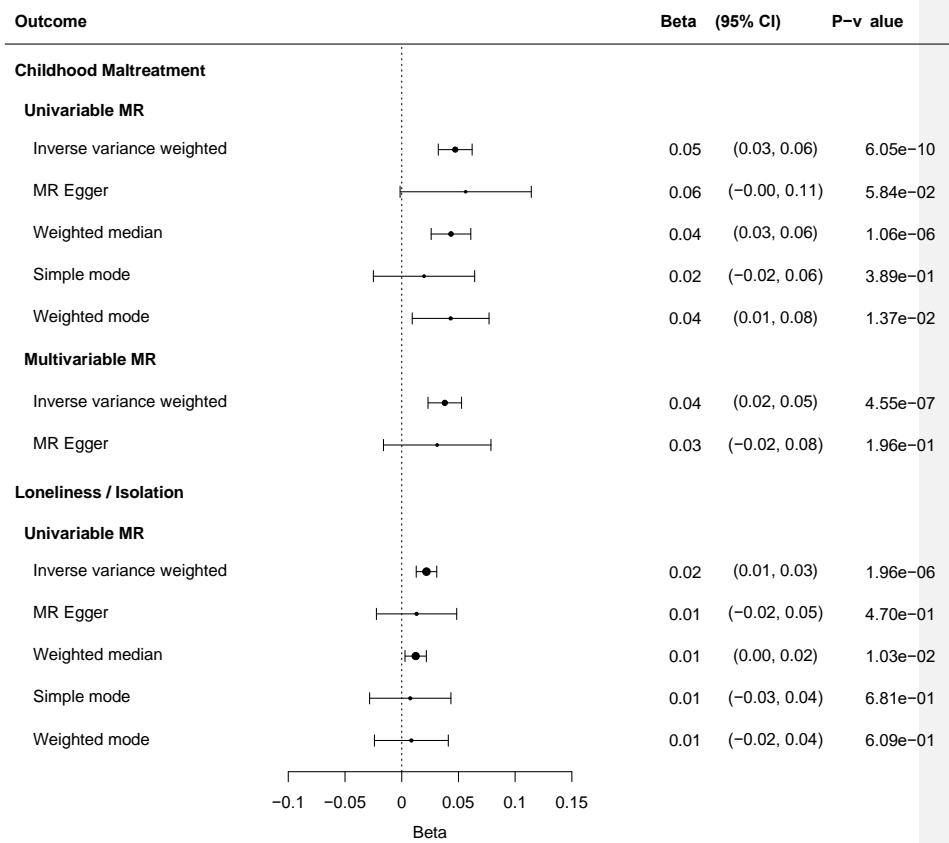
*Univariable and Multivariable Mendelian Randomisation Estimates for the Effect of Genetically Predicted Childhood Maltreatment and Loneliness/Isolation on Schizophrenia Risk*



*Note.* Error bars represent 95% confidence intervals. MR = Mendelian randomisation; CI = confidence interval. Confidence intervals of the MR-Egger estimate for childhood maltreatment are wider, consistent with the lower statistical power of this method (Slob & Burgess, 2020).

**Figure 4**

*Univariable and Multivariable Mendelian Randomisation Estimates for the Effect of Genetically Predicted Schizophrenia on Childhood Maltreatment and Loneliness/Isolation Risk*



*Note.* Error bars represent 95% confidence intervals. MR = Mendelian randomisation; CI = confidence interval.