

Validation of a Prediction Rule for Mortality in Congenital Diaphragmatic Hernia

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abstract

BACKGROUND: Congenital diaphragmatic hernia (CDH) is a rare congenital anomaly with a mortality of ~27%. The Congenital Diaphragmatic Hernia Study Group (CDHSG) developed a simple postnatal clinical prediction rule to predict mortality in newborns with CDH. Our aim for this study is to externally validate the CDHSG rule in the European population and to improve its prediction of mortality by adding prenatal variables.

METHODS: We performed a European multicenter retrospective cohort study and included all newborns diagnosed with unilateral CDH who were born between 2008 and 2015. Newborns born from November 2011 onward were included for the external validation of the rule ($n = 343$). To improve the prediction rule, we included all patients born between 2008 and 2015 ($n = 620$) with prenatally diagnosed CDH and collected pre- and postnatal variables. We build a logistic regression model and performed bootstrap resampling and computed calibration plots.

RESULTS: With our validation data set, the CDHSG rule had an area under the curve of 79.0%, revealing a fair predictive performance. For the new prediction rule, prenatal herniation of the liver was added, and absent 5-minute Apgar score was taken out. The new prediction rule revealed good calibration, and with an area under the curve of 84.6%, it had good discriminative abilities.

CONCLUSIONS: In this study, we externally validated the CDHSG rule for the European population, which revealed fair predictive performance. The modified rule, with prenatal liver herniation as an additional variable, appears to further improve the model's ability to predict mortality in a population of patients with prenatally diagnosed CDH.



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WHAT'S KNOWN ON THIS SUBJECT: Mortality rate in newborns with congenital diaphragmatic hernia is substantial. There are pre- and postnatal predictors for mortality, but many have considerable disadvantages. The Congenital Diaphragmatic Hernia Study Group developed a simple clinical prediction rule to predict mortality early in life.

WHAT THIS STUDY ADDS: We validated the Congenital Diaphragmatic Hernia Study Group prediction rule in the European population. The rule revealed fair predictive performance. The modified rule, with prenatal liver herniation as an additional variable, improved the model's ability to predict mortality in patients with prenatally diagnosed congenital diaphragmatic hernia.

To cite: Cochijs-den Otter SCM, Erdem Ö, van Rosmalen J, et al. Validation of a Prediction Rule for Mortality in Congenital Diaphragmatic Hernia. *Pediatrics*. 2020;145(4):e20192379

Congenital diaphragmatic hernia (CDH) is a severe developmental defect of the diaphragm that causes lung hypoplasia and pulmonary hypertension (PH), leading to a mortality of 27% in live-born patients.¹ Identification of risk factors that prognosticate outcome in patients with CDH is essential to accurately counsel parents and to compare patient populations and management strategies.

Prenatally, outcomes are predicted by using the observed-to-expected lung-to-head ratio (O/E LHR), MRI calculations of lung volumes, and the position of the liver and stomach.²⁻⁷ These prenatal parameters can be used to predict lung hypoplasia but do not seem to reliably predict PH.^{8,9}

For the postnatal prediction of survival, there are several prediction models and variables, such as the Score for Neonatal Acute Physiology-Perinatal Extension II score and oxygenation index. However, many are based on relatively small groups of patients, are difficult to apply, or have not been externally validated.¹⁰⁻¹⁴

Brindle et al,¹⁵ and the Congenital Diaphragmatic Hernia Study Group (CDHSG) have developed a simple early clinical prediction rule in a large cohort of patients to identify low (<10%), intermediate (~20%), and high risk (~50%) of death in the postnatal period. This prediction model is based on birth weight, the 5-minute Apgar score, severe PH, and the presence of cardiac and chromosomal anomalies. Validation of the prediction rule revealed reasonable discrimination between groups.^{15,16}

This postnatal model has been favorably compared with prenatal predictors.¹⁷ However, there is potential value in combining post- and prenatal risk factors within a single prediction model. Pre- and postnatal predictors have only been integrated in one prediction model in

a small group of patients from a single center.¹⁸ Our aim for this study was to externally validate the CDHSG clinical prediction rule in a European population and incorporate additional prenatal variables to further improve the rule.

METHODS

The data were collected from 4 high-volume CDH centers treating ≥ 10 patients with CDH per year.¹⁹ These centers are part of the CDH Euro Consortium: Erasmus University Medical Center, Rotterdam, Netherlands; Radboud University Medical Center Amalia Children's Hospital, Nijmegen, Netherlands; University Hospital Mannheim, Mannheim, Germany; and Bambino Gesù Children's Hospital, Rome, Italy. The CDH Euro Consortium is a voluntary collaboration of European institutions that work together in research. This collaborative group also developed the CDH Euro Consortium management guidelines that are implemented in all participating centers.^{1,20} Institutional review board approval was obtained from the Medical Ethics Committee, Erasmus University Medical Center in Rotterdam (MEC2016-109).

For the external validation of the CDHSG prediction rule, patients born before November 2011 were excluded because these patients were included in the CDHSG database and were used for the development of the original CDHSG prediction rule.¹⁵ We included all live-born infants with CDH who were born between November 2011 and November 2015. We reviewed the data of these patients from the local CDHSG database and added missing data from the medical files if available. The collected data were in accordance with the definitions used by Brindle et al¹⁵: low birth weight (<1500 g); an Apgar score <7 at 5 minutes or the absence of an Apgar score; severe

PH, defined as right-to-left shunt or estimated supra-systemic pulmonary pressures on the first echocardiogram; chromosomal anomalies, defined as any abnormalities in the chromosomal array; major cardiac anomalies, classified as all anomalies other than a patent foramen ovale; patent ductus arteriosus; an atrial septum defect; and a ventricular septum defect.

The data of each patient were entered in the CDHSG prediction rule to calculate a total CDH risk score, ranging from 0 to 8 (Table 1). This score was used to stratify the patients into 1 of the 3 risk groups; low (0), intermediate,^{1,2} and high risk.³⁻⁸

For the implementation of prenatal variables in the CDHSG prediction rule, we included all live-born infants with prenatally diagnosed CDH who were born between 2008 and 2015. The predictors in the CDHSG prediction rule were reviewed. Most of the variables were used as binary variables. However, to further improve the model, birth weight was also tested as a continuous variable, and low Apgar score was defined as <5 at 5 minutes or <7 at 5 minutes. Missing Apgar score was left out because one of the centers never calculates an Apgar score for patients with CDH. Also, after discussion with an expert group consisting of pediatric intensivists, neonatologists, and prenatal specialists across participating centers, we decided that the variable chromosomal anomalies should always be in the model because of its major significance in the decision to start and continue treatment.

Additionally, candidate pre- and postnatal predictors were selected by the expert group. The first measured O/E LHR after 18 weeks' gestation was included as a continuous variable. The presence of intrathoracic liver herniation on the last prenatal ultrasound was used as a binary variable. Also, the side of the

TABLE 1 CDHSG Prediction Rule and the New Model

	Description	Value
Original CDHSG prediction equation	$1/(1 + \exp(2.65 - \log(2.634) \times (\text{low birth wt}) - \log(2.718) \times (\text{low 5-min Apgar score} < 7) - \log(4.678) \times (\text{missing 5-min Apgar score}) - \log(4.073) \times (\text{severe PH}) - \log(5.22) \times (\text{MCAs}) - \log(3.928) \times (\text{chromosomal anomaly}))$	%
Final CDHSG prediction rule ¹⁵	Low birth wt (<1500 g) Low 5-min Apgar score (<7) Missing 5-min Apgar score Severe PH MCAs Chromosomal anomaly Total CDHSG score (sum values)	1 1 2 2 2 1 0–8
New prediction model with additional prenatal variable	$1/(1 + \exp(-0.6735 + 0.0013 \times (\text{birth wt [g]}) - 1.7150 \times (\text{low 5-min Apgar score} < 7) - 1.4871 \times (\text{severe PH}) - 0.9471 \times (\text{MCAs}) - 0.8754 \times (\text{chromosomal anomaly}) - 0.7235 \times (\text{intrathoracic liver herniation on prenatal ultrasound}))$	%

MCA, major cardiac anomaly.

hernia, fetal endotracheal occlusion, the presence of polyhydramnios,²¹ gestational age at diagnosis, and gestational age at birth were selected.

To describe the baseline characteristics of the patients with CDH, medians and interquartile ranges (IQRs) were used for continuous variables, and percentages were used for categorical variables. Comparisons between baseline characteristics and death before discharge were made by using the χ^2 test for categorical variables and the Mann–Whitney *U* test for continuous variables. Comparisons between centers were made by using Kruskal–Wallis and χ^2 tests.

For the external validation of the CDHSG prediction rule, multiple imputation was performed for missing data (Table 2), creating 100 databases by using fully conditional specification. Because the available data between centers were heterogeneous, we used “center” as a covariate in the multiple imputation. Then the CDHSG prediction score was calculated for each individual by using the final prediction rule, as used by Brindle et al,¹⁵ as well as the original equation, which was used to develop the CDHSG prediction rule (Table 1). The predictive performance was

assessed by using calibration plots and the *c*-statistic (ie, the area under the receiver operating characteristic [ROC] curve). Also, the predicted outcome of the final equation was compared with the observed outcome in the study cohort from the pooled database.

For the new model, predictors were tested by using univariate analysis, assessing if a variable was associated with increased mortality. We corrected for center. The selected variables were put into a multivariable logistic regression model by using the stepwise backward method. In every step, the variable with the highest *P* value was excluded if *P* > .1, and this was repeated until all variables included in the model had *P* < .1. The model was evaluated with a calibration plot to assess the discriminatory abilities of the model, followed by bootstrapping to correct for the optimism of the model. We then calculated the predicted risk per patient and plotted the ROC curves to determine cutoff values of the predicted risk for 3 risk groups: low, intermediate, and high risk. SPSS Statistics version 24 (IBM SPSS Statistics, IBM Corporation, Armonk, NY) and R version 3.6.1, with the packages *rms* and *mice*, were used for the statistical analyses.

RESULTS

A total of 753 patients were diagnosed with CDH between January 1, 2008, and December 31, 2015. Eight patients were excluded because there were no patient characteristics available. Fourteen pregnancies resulted in an intrauterine fetal demise. Three hundred forty-three patients were born between 2011 and 2015, and their data were used for the validation of the original prediction rule. Six hundred twenty patients were included to develop the new rule. In 111 patients, the diagnosis was not prenatally known, and therefore they were excluded for the new rule. This postnatally diagnosed group had a mortality of 9%.

Baseline characteristics of both patient groups are shown in Table 2. In 70.3% of the patients in the cohort used for the validation, the first echocardiogram was performed within the first 24 hours of life. The overall mortality was 18%. In the group used for the new rule, 76.9% of the patients had their first echocardiogram performed within the first 24 hours of life. Their overall mortality was 23%. In both groups, the baseline characteristics of the patients who survived were significantly different from those of patients who died, except for sex

TABLE 2 Patient Characteristics for the Validation of the CDHSG Rule and the New Model

	Patients for Validation (n = 343)		Survivors (n = 280)		Nonsurvivors (n = 63)		P	Patients, New Model (n = 620)		No. Missing Data, n (%)		Survivors (n = 478)		Nonsurvivors (n = 142)		P	
Male sex	183 (56%)	0	156 (56%)	37 (59%)	663	366 (59%)	0	288 (60%)	78 (55%)	258							
First measured O/E LHR	40.0 (IQR 33.4–49.5)	99 (29)	41.5 (IQR 34.0–50.1)	37 (IQR 25.8–47.1)	<.05	39.8 (IQR 32.3–49.5)	127 (20)	41.3 (IQR 33.9–50.5)	34.9 (IQR 27.0–42.6)	<.05							
Intrathoracic liver herniation	176 (51%)	44 (13)	127 (45%)	49 (78%)	<.05	335 (54%)	18 (3)	234 (49%)	101 (71%)	<.05							
Prenatal surgery	25 (7%)	0	17 (6%)	8 (13%)	.068	62 (10%)	0	34 (7%)	28 (20%)	<.05							
Gestational age, wk	38.0 (IQR 37.0–38.6)	3 (<1)	38.0 (IQR 37.1–38.6)	37.1 (IQR 35.6–38.6)	<.05	38.0 (IQR 36.6–38.4)	4 (1)	38.0 (IQR 37.0–38.6)	37.0 (IQR 35.0–38.8)	<.05							
Left-sided defect	292 (85%)	0	242 (86%)	50 (79%)	.135	536 (87%)	4 (1)	428 (90%)	108 (76%)	<.05							
Apgar score at 5 min	8 (IQR 7–9)	58 (17)	8 (IQR 7–9)	7 (IQR 6–8)	<.05	8 (IQR 7–8)	131 (21)	8 (IQR 7–9)	7 (IQR 6–8)	<.05							
Apgar score <7 at 5 min	56 (16%)	58 (17)	34 (12%)	22 (35%)	<.05	86 (14%)	131 (21)	39 (8%)	47 (33%)	<.05							
Birth wt, kg	3.00 (IQR 2.67–3.33)	1 (<1)	3.00 (IQR 2.70–3.35)	2.80 (IQR 2.15–3.20)	<.05	2.97 (IQR 2.56–3.25)	3 (<1)	3.00 (IQR 2.69–3.30)	2.63 (IQR 2.04–3.00)	<.05							
Chromosomal anomalies	19 (6%)	0	11 (4%)	8 (13%)	<.05	23 (4%)	5 (1)	13 (3%)	10 (7%)	<.05							
PH	79 (23%)	66 (19)	47 (17%)	32 (51%)	<.05	169 (27%)	123 (20)	99 (21%)	70 (49%)	<.05							
ECMO treatment	110 (32%)	2 (<1)	74 (26%)	36 (57%)	<.05	190 (31%)	0	122 (26%)	68 (48%)	<.05							
Survival	280 (82%)	0	—	—	—	478 (77%)	0	—	—	—							

Patient characteristics were specified for survivors and nonsurvivors. Patients were used for the validation of the CDHSG rule on the left side, and patients were used for the new model on the right side. —, not applicable.

TABLE 3 Patient Characteristics Specified per Center for the New Model (n = 620)

Center	First Measured O/E LHR, % (IQR)	Prenatal Ultrasound		Prenatal Surgery	Left-Sided Hernia	5-min Apgar Score <7	Birth Wt, kg (IQR)	Major Cardiac Anomalies		Chromosomal Anomalies	Severe PH	ECMO	Survival
		Intrathoracic Liver Herniation on	Prenatal Ultrasound					Major Cardiac Anomalies	Chromosomal Anomalies				
Rotterdam	43.2 (34.1–53.4)	32%	5%	86%	18%	3.0 (2.5–3.2)	4%	2%	16%	30%	72%		
Nijmegen	42.3 (36.6–50.4)	40%	11%	84%	19%	2.9 (2.6–3.3)	2%	7%	58%	44%	66%		
Mannheim	37.8 (30.6–46.1)	68%	13%	88%	15%	2.9 (2.6–3.3)	5%	5%	21%	37%	83%		
Rome	45.1 (35.6–58.6)	45%	4%	84%	0% ^a	3.0 (2.5–3.3)	8%	0%	45%	0%	70%		
P	<.01	<.01	<.01	.935	.151	.963	.382	.087	<.01	<.01	<.01		

^a The Apgar score is never calculated in this center.

(Table 2). These characteristics also differed significantly between centers, as presented in Table 3.

The outcome of the CDHSG prediction rule after multiple imputations is shown in Table 4. Forty-six percent of the patients were grouped in the low-risk group (score 0), with an observed mortality of 4%, and 38% of patients were grouped in the intermediate group (score 1–2), with a mortality of 22%. The high-risk group (score 3–8) was smaller, containing 16% of the patients, with a mortality of 66%. The discrimination of the model was moderately strong, with a c-statistic of 0.784 for the original equation and 0.790 for the final CDHSG prediction rule.

Subsequently, to develop a new rule, the original prediction rule was modified. First, logistic regression was performed within the large data set by using a backward elimination algorithm. Missing data were imputed. O/E LHR, side of the hernia, gestational age at birth, fetal endotracheal occlusion, polyhydramnios, Apgar score <5 at 5 minutes, and gestational age at diagnosis were excluded from the model with backward elimination. Although chromosomal anomalies had a *P* value >.1, we forced it into the model (Table 5). The new model contains birth weight as a continuous variable and intrathoracic herniation of the liver, major cardiac anomalies, chromosomal anomalies, Apgar score <7 at 5 minutes, and severe PH as binary variables (Table 1). An evaluation of the model in a calibration plot revealed good discrimination of the model, with a c-statistic of 0.859. When correcting

TABLE 5 Odds Ratios for Mortality for Variables in the New Model

Variable	Adjusted OR	95% Confidence Interval
Intercept	1.9611	0.5570–6.9048
Birth wt, g	0.9987	0.9983–0.9991
Intrathoracic liver herniation	2.0616	1.2300–3.4555
MCAs	2.5781	0.9631–6.9020
Chromosomal anomalies	2.3998	0.8277–6.9579
Severe PH	4.4242	2.6159–7.4826
Apgar score <7	5.5567	3.0719–10.0513

MCA, major cardiac anomaly; OR, odds ratio.

for the optimism of the model, estimated ~1.4%, the c-statistic is 0.846. In Supplemental Fig 1, the ROC curve of the new model is shown. We then stratified the patients into 1 of the 3 groups: low, intermediate, and high risk of mortality. When using <10% (mild), 10% to 50% (moderate), and >50% (severe) risk of mortality as cutoff points, the cutoff between the mild group and the moderate group revealed a sensitivity of 90.8% and a specificity of 55.4%, whereas the cutoff between the moderate and the severe group revealed a sensitivity of 49.3% and a specificity of 93.5%.

The disease severity, by using the rules per center, is presented in Supplemental Tables 6 and 7.

DISCUSSION

In this study, we externally validated the CDHSG rule in the European population. We found that the rule had fair discrimination, but also room for optimization, compared with the internal validation of Brindle et al.¹⁵ Bent et al¹⁶ also validated the rule in a large group of patients with CDH born in California and found an underestimation of mortality in the patients with a score of 1. We did not find this in our population. This might

be explained by the difference in health care systems in Europe and the United States. In Europe, centralized care is more common and many patients with CDH are born in high-volume centers. It is increasingly recognized that centralized care improves outcome in these patients.¹⁹ This might also explain the lowest mortality in patients born in the largest center of our study. Furthermore, in Europe, many CDH centers collaborate in the CDH Euro Consortium, which has developed a standardized treatment protocol, increasing survival from 67% to 88%.^{20,22}

Although it seems valid to use the model, some variables were not useful or difficult to apply in our population. In 1 of the 4 centers, Apgar scores were never measured because the medical team felt it was not a useful tool in this patient group. Patients with CDH will have a lower Apgar score because they are intubated directly after birth.²⁰ Brindle et al¹⁵ theorized that the absence of an Apgar score implies an infant is sicker, but this was not applicable to our cohort. Also, the measurement of PH on echocardiography is not standardized, and different definitions for PH are being used. Brindle et al¹⁵ used right-

TABLE 4 CDHSG Prediction Rule: Predicted and Observed Mortality Risk After Multiple Imputation

	CDHSG Score						
	0 (n = 157.0)	1 (n = 34.5)	2 (n = 96.9)	3 (n = 34.3)	4 (n = 17.8)	5 (n = 1.3)	6 (n = 1.3)
Predicted mortality, %	6.6	17.1	24.1	45.5	60.1	87.9	87.5
Observed mortality, %	4.0	18.5	22.9	42.7	63.6	87.5	100

to-left shunting or estimated suprasystemic pulmonary pressures. In Europe, the presence of PH is often defined as pulmonary pressures higher than two-thirds of the systemic pressures.^{20,23,24} Furthermore, the timing of the measurement differs between centers. Brindle et al¹⁵ used the earliest echocardiogram in the model, whereas Bent et al¹⁶ used PH at discharge. Presumably, the incidence of severe PH was underestimated in the study by Bent et al¹⁶ because many patients with PH would have already died, and in others, pulmonary pressures would have decreased.²⁵ The registered incidence of PH in nonsurvivors is only 33.5% in the study by Bent et al¹⁶, whereas it is >50% in our cohort and >60% in the CDHSG population, supporting this assumption.^{15,16}

To improve the power of the original prediction rule, combining pre- and postnatal variables is presumably superior. Prenatal variables have been found to adequately predict lung hypoplasia and the need for extracorporeal membrane oxygenation (ECMO) but are less reliable as a marker for PH.^{8,9} To predict mortality, postnatal variables are still essential. Different prenatal variables were tested in the model, and eventually, only the position of the liver was a significant variable. Surprisingly, the O/E LHR was not of additional value to the model, whereas in earlier studies, the O/E LHR did have a role in predicting survival. It is a more reliable prenatal predictor than the lung-to-head-ratio because it is a stable variable during pregnancy.^{6,26} However, different O/E LHR measurement techniques are being used between centers, and there is a learning curve in the examination of the O/E LHR.^{27,28} In 2 of the centers, the tracing method was always used, whereas in the others, also the longest-axis-diameter method was used. The longest-axis-diameter method overestimates the

O/E LHR up to 34% and has a larger interobserver variability.²⁹⁻³¹ In addition, the O/E LHR can be calculated with multiple calculators (ie, www.totaltrial.eu or www.perinatology.com), resulting in different ratios. In our study, the method of calculation varied. Measuring lung volumes on MRI holds promise.³² However, in many centers, it is not possible to use MRI for this purpose because of costs and lack of availability. For this study, analysis of fetal lung volume resulted in too many missing data. Another prenatal predictor is stomach position.⁷ This measure, however, is not implemented in standard prenatal care and could not therefore be analyzed for this study. More consistent measuring techniques and reporting would probably make these different prenatal variables more suitable for use in outcome prediction.

For the modified model, we also made some changes to the original variables. "Missing Apgar score" was taken out. On the other hand, we kept chromosomal anomalies in the model, although its association with mortality was not significant ($P > .1$). Possibly, some patients had a clinically insignificant abnormality on array; however, because these data are retrospective, it was not possible to reliably select only clinically relevant cases. Therefore, we choose to include all patients with abnormalities on chromosomal array. CDH is associated with numerous chromosomal anomalies, and often the associated anomaly has a major impact on mortality.^{33,34}

The strength of our study is the large population of patients with CDH and the amount of prenatal data available. Often, the implementation of prenatal data in a prediction model is difficult because of the amount of missing data.¹⁵ Also, when using a voluntary database, or a database based on coded diagnoses, the accuracy of the data are difficult to interpret. We were able to go back to the original

patient files when needed. However, we did have some missing data, as shown in Table 2, which we corrected for using multiple imputation, a statistical tool often used in this setting. However, it is possible that some predictors were not significant in the model because of the large amount of missing data, such as the presence of polyhydramnios and gestational age at birth. Furthermore, because all 4 centers are part of the CDH Euro Consortium, postnatal clinical management is similar in all centers, increasing the reliability of an early prediction model, but it also potentially limits the generalizability of our study and may contribute to an optimistic assessment of model performance. Therefore, our new model needs additional external validation in a more heterogeneous group like the patients in the CDHSG database to prove its generalizability. Also, because prenatal data are necessary for the new rule, other settings with imperfect prenatal care provision may not benefit as much from this new model. The original CDHSG model is easy to apply at the bedside. For our model, a more complicated calculation is necessary.

In a population with a rare congenital defect with high mortality and morbidity, it is important to reliably predict outcomes. Prenatally, this can guide parents and clinicians in decisions regarding perinatal management and in the referral to high-volume centers for the delivery in areas with a low-density population. Postnatally, adequate prediction can help parents to better understand the course of their child's illness. In addition, it can also be used for standardized reporting and benchmarking between centers. A good postnatal prediction model can potentially improve care for specific groups of patients with CDH. This model can act as a practical tool when stratifying patients into risk groups such as mild, moderate, and severe. The mild group of patients, identified

with a highly sensitive threshold, could receive less-aggressive treatment, such as a spontaneous breathing trial at birth.³⁵ On the other hand, the severe group, identified with a highly specific threshold, could potentially benefit most from more aggressive experimental therapies, and the true benefit of these therapies could be detected earlier because there would be no dilution of effect due to the inclusion of patients with lower risk. For all these reasons, there is a need for a reliable prediction model that can be applied on the first day of life.³⁶ The power of a prediction model is dependent on the availability of the information needed for the calculation. Prenatal data are not readily available in some areas of the world. However, the only prenatal parameter significant in our

model, the prenatal position of the liver on ultrasound, is reasonably easy to evaluate.

CONCLUSIONS

We have successfully validated the CDHSG prediction rule within a European population. We also developed a modification of the original rule, implementing prenatal variables, with apparent improvement of the predictability of mortality. Standardization of the measurements of prenatal variables, such as the O/E LHR, and the postnatal variable PH could potentially increase their predicting value and further improve these models. Validation of this modified rule is needed to evaluate its generalizability.

ACKNOWLEDGMENTS

We thank Dr R. Crisafulli for the data collection. We thank Prof K. Allegaert and Dr R. de Jonge for their useful contribution to the article.

ABBREVIATIONS

CDH: congenital diaphragmatic hernia
CDHSG: Congenital Diaphragmatic Hernia Study Group
ECMO: extracorporeal membrane oxygenation
IQR: interquartile range
O/E LHR: observed-to-expected lung-to-head ratio
PH: pulmonary hypertension
ROC: receiver operating characteristic

Dr Falk collected data and critically revised the manuscript; Dr Brindle contributed to the statistical concept of the model, gave insight into the original data and model, and revised the manuscript; Dr Tibboel conceptualized and designed the study, took part in the expert group for the development of the new model, and revised the manuscript; and all authors approved the final manuscript as submitted and agree to be accountable for all aspects of the work.

DOI: <https://doi.org/10.1542/peds.2019-2379>

Accepted for publication Jan 27, 2020

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PEDIATRICS (ISSN Numbers: Print, 0031-4005; Online, 1098-4275).

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FINANCIAL DISCLOSURE: The authors have indicated they have no financial relationships relevant to this article to disclose.

FUNDING: No external funding.

POTENTIAL CONFLICT OF INTEREST: The authors have indicated they have no potential conflicts of interest to disclose.

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Pediatrics 2020;145;

DOI: 10.1542/peds.2019-2379 originally published online March 5, 2020;

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