

EPIDEMIOLOGY

In-person schooling and associated COVID-19 risk in the United States over spring semester 2021

Kirsten E. Wiens¹, Claire P. Smith¹, Elena Badillo-Goicoechea², Kyra H. Grantz¹,
M. Kate Grabowski^{1,3}, Andrew S. Azman^{1,4}, Elizabeth A. Stuart^{2,5,6}, Justin Lessler^{1,7,8*}

Because of the importance of schools to childhood development, the relationship between in-person schooling and COVID-19 risk has been one of the most important questions of this pandemic. Previous work in the United States during winter 2020–2021 showed that in-person schooling carried some risk for household members and that mitigation measures reduced this risk. Schooling and the COVID-19 landscape changed radically over spring semester 2021. Here, we use data from a massive online survey to characterize changes in in-person schooling behavior and associated risks over that period. We find increases in in-person schooling and reductions in mitigations over time. In-person schooling is associated with increased reporting of COVID-19 outcomes even among vaccinated individuals (although the absolute risk among the vaccinated is greatly reduced). Vaccinated teachers working outside the home were less likely to report COVID-19–related outcomes than unvaccinated teachers working exclusively from home. Adequate mitigation measures appear to eliminate the excess risk associated with in-person schooling.

INTRODUCTION

The role of children and in-person schooling in severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2) transmission continues to be a contentious issue. Policies regarding in-person schooling have varied markedly across school districts in the United States, with a heterogeneous mix of in-person and remote learning as well as varying approaches to mitigation (1). Over the spring semester of the 2020–2021 school year, many school districts made major updates to their approach to in-person schooling as the winter wave of the coronavirus disease 2019 (COVID-19) pandemic receded. Unfortunately, resurgences related to the Delta and Omicron variants in the 2021–2022 school year have meant that, at the time of writing, COVID-19 remains a major health threat in the United States and worldwide (2, 3). However, in light of increased vaccine availability and a broad consensus on the benefits of in-person schooling, the vast majority of school districts in the United States conducted in-person classes in fall semester 2021 despite this surge in cases (4). Hence, understanding the risks associated with in-person schooling and how best to control them remains an important area of research.

A challenge in quantifying the risk of in-person schooling has been limited information on the degree to which children, who mostly have mild or asymptomatic infections, can infect teachers and family members (5). In an effort to address this issue, we previously analyzed data collected by the nationwide U.S. COVID-19 Trends and Impact Survey (U.S. CTIS) between 24 November 2020 and

10 February 2021 (6). This survey is administered by the Delphi Group in partnership with Facebook and includes questions on demographics, symptoms related to COVID-19, positive SARS-CoV-2 test results, and schooling for any children in the household (7, 8). We showed that individuals in a home with a child engaged in in-person schooling were at significantly higher risk of developing COVID-19–related outcomes (9). We also found that this risk decreased with increasing numbers of school-based mitigation measures, with no additional risk—as compared to no in-person schooling—associated with children attending schools with seven or more mitigation measures (9).

Between the time of this original study and the end of spring semester 2021 in June, there were major changes in both the pandemic situation and schooling policies in most U.S. schools. These include rising vaccination rates and the rise of the Alpha (and later the Delta) variant in the United States, both of which had major impacts on the dynamics of the COVID-19 epidemic (10). The proportion of the U.S. adult population who received at least one COVID-19 vaccine dose increased from about 7% at the end of January to 53% by mid-June (11). Beginning in mid-February 2021, the Alpha variant, which is roughly 50% more transmissible than the previously dominant SARS-CoV-2 strains (12), spread rapidly throughout the country (3, 13). In June, the Delta variant, which is about 60% more transmissible than Alpha (14), began to dominate (3, 13). It was unclear how these factors modified the risk posed by in-person schooling.

In this study, we expand our previous analysis (9) to include data from the U.S. CTIS for the entire spring semester 2021 (which includes 4 weeks from the previous study), with the goals of (i) characterizing how rates of in-person schooling and implementation of school-based mitigation measures changed over the course of the semester, (ii) understanding whether and how vaccination status and Alpha/Delta variant prevalence modified the association between household COVID-19 risk and living with a child in in-person schooling, and (iii) identifying other temporal trends in the relationship between in-person schooling and the risk of household members reporting COVID-19–related outcomes.

¹Department of Epidemiology, Johns Hopkins Bloomberg School of Public Health, Baltimore, MD, USA. ²Department of Mental Health, Johns Hopkins Bloomberg School of Public Health, Baltimore, MD, USA. ³Department of Pathology, Johns Hopkins University School of Medicine, Baltimore, MD, USA. ⁴Institute of Global Health, Faculty of Medicine, University of Geneva, Geneva, Switzerland ⁵Department of Biostatistics, Johns Hopkins Bloomberg School of Public Health, Baltimore, MD, USA. ⁶Department of Health Policy and Management, Johns Hopkins Bloomberg School of Public Health, Baltimore, MD, USA. ⁷Department of Epidemiology, University of North Carolina at Chapel Hill, Chapel Hill, NC, USA. ⁸Carolina Population Center, University of North Carolina at Chapel Hill, Chapel Hill, NC, USA.

*Corresponding author. Email: jlessler@unc.edu

RESULTS

We analyzed data from 1,082,773 respondents living with school-aged children in 50 U.S. states and Washington, DC that were collected through the Delphi Group at Carnegie Mellon University U.S. CTIS from 12 January 2021 to 12 June 2021 (table S1). Although the total number of respondents decreased over the study period, patterns in the relative number of respondents by county remained consistent (Fig. 1A).

Overall, 59.4% of respondents living with school-aged children reported having at least one child in their household attending in-person schooling. The proportion of respondents living with school-aged children reporting any in-person schooling increased from 47.0% during the week of 12 January 2021 to 65.3% during the week ending 12 June 2021 (Figs. 1B and 2A). Of those reporting any in-person schooling, 74.0% reported full-time in-person instruction, increasing from 69.0 to 82.1% over the study period (Fig. 1C).

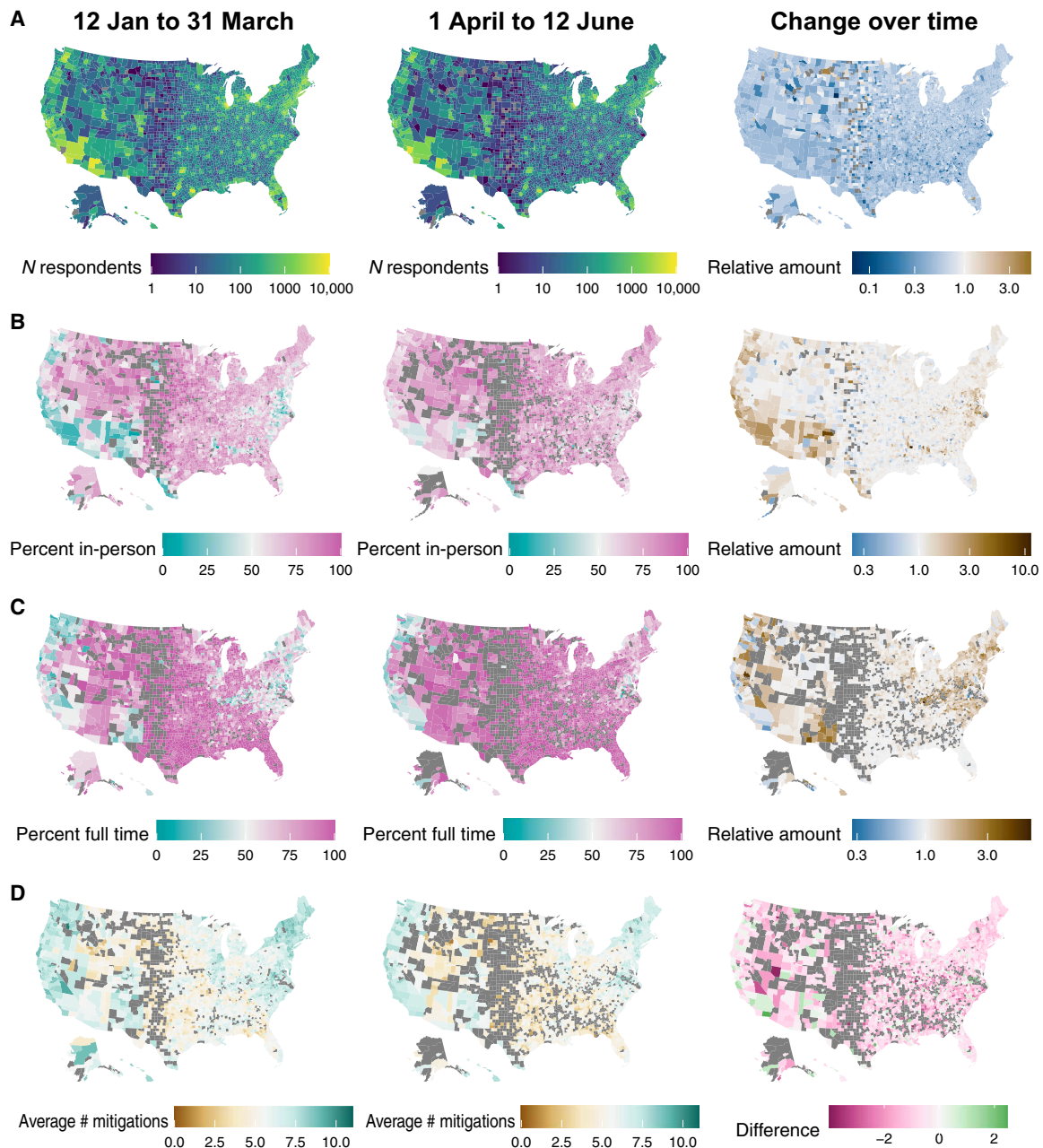


Fig. 1. Changes over time in in-person schooling by county. Distribution of survey responses from 12 January to 31 March (left column), 1 April to 12 June (center column), and change over time (right column). Results are shown for (A) number (n) of survey respondents reporting ≥ 1 school-aged child in the household, (B) percent reporting in-person schooling, (C) percent of respondents with in-person schooling reporting full-time in-person instruction, and (D) average number of school-based mitigation measures. “Relative amount” in the right column indicates values in 1 April to 12 June (center column) divided by values in 12 January to 31 March (left column). Gray indicates county/periods where fewer than 10 respondents reported in-person schooling.

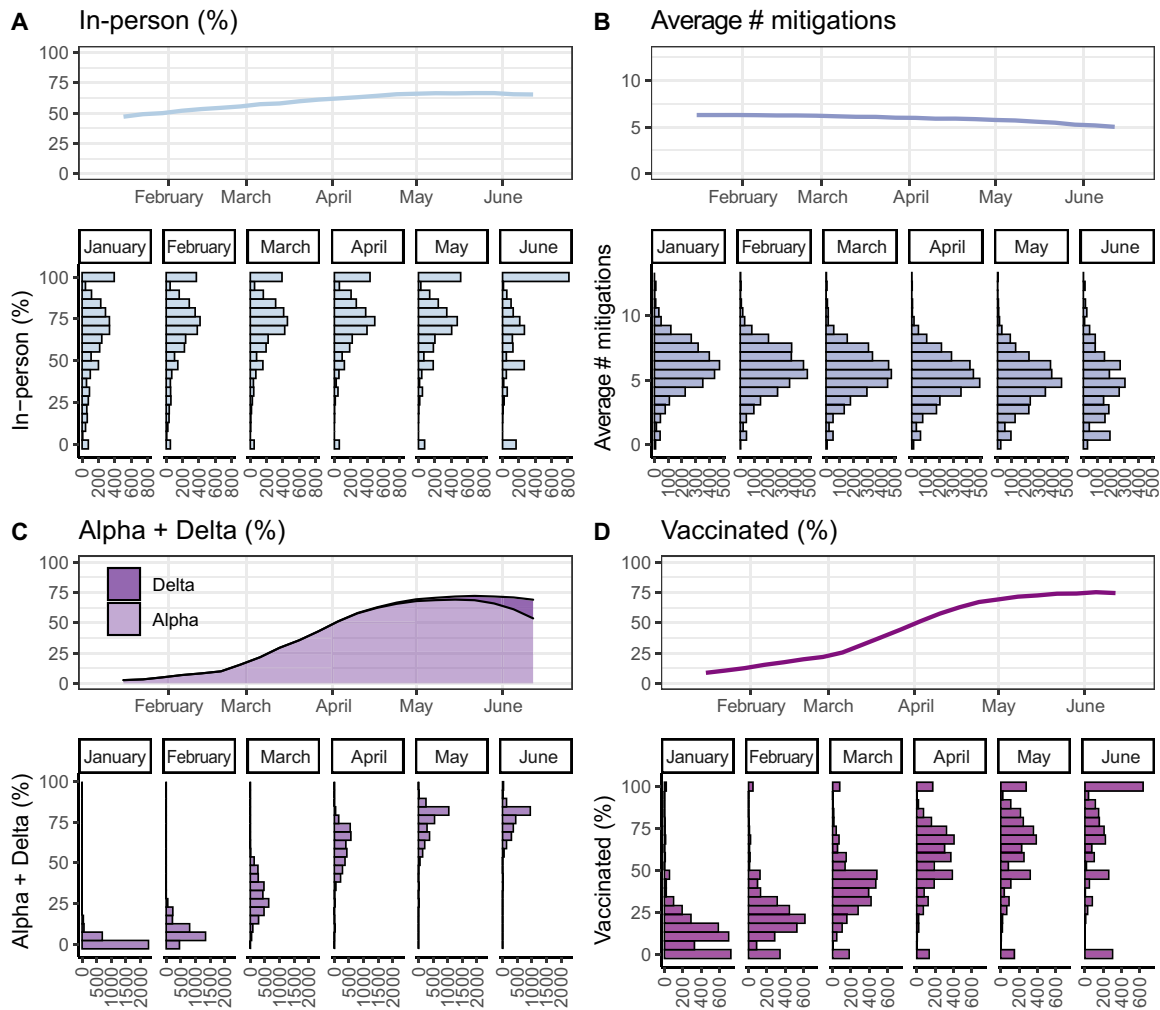


Fig. 2. Concurrent changes in in-person schooling, vaccination, and variant prevalence. Changes over time between 12 January and 12 June in (A) percent of respondents living with school-aged children reporting any in-person schooling, (B) average number of school-based mitigation measures, (C) smoothed percent of Global Initiative on Sharing All Influenza Data (GISAID) SARS-CoV-2-sequenced isolates that were Alpha or Delta, and (D) percent of respondents living with school-aged children reporting having received at least one COVID-19 vaccine dose. Top panel line plots show national averages by week weighted using U.S. CTIS survey weights (A and B), state population (C), and county population (D). Bottom panel histograms show the number of county months with the indicated percentages (A, C, and D) or numbers (B). Note that the number of respondents decreased over time (Fig. 1D), which may contribute to the increasing number of zero values in the histograms in June.

Of those respondents living with a child attending school in-person, 93.1% reported at least one mitigation measure in place, 72.2% reported at least four, and 46.9% reported at least seven. The average number of mitigation measures decreased from 6.3 to 5.0 over the study period (Figs. 1D and 2B). The biggest changes in in-person schooling behavior occurred in places with the least in-person schooling at the start of the study period (Fig. 1, B to D), although the geographic regions with the most cautious approaches to in person instruction remained largely the same.

Since the seventh wave of the survey, starting on 12 January 2021, the U.S. CTIS has included questions on the respondent’s vaccination status and number of doses received (7, 8). Overall, 43.1% of 1,082,773 respondents living with school-aged children reported having received at least one COVID-19 vaccine dose compared with 50.8% among all 5,273,116 survey respondents. Although vaccination among those with school-aged children was lower on average (fig. S1A), the magnitude of these differences varied by county (fig. S1B).

Vaccination among survey respondents living with school-aged children was strongly correlated with rates reported by the U.S. Centers for Disease Control and Prevention (CDC) [Pearson’s correlation coefficient ($r = 0.83$) (15), but rates were higher among survey respondents in most counties (particularly those in New Mexico, South Dakota, and Nebraska) (fig. S1, C and D).

While in-person schooling increased over the study period, vaccination rates and the prevalence of Alpha and Delta variants increased to a greater extent (Fig. 2 and movie S1). Because of the low proportion of cases from the Delta variant over the study period, we analyze the Alpha and Delta variants together throughout this manuscript. The prevalence of Alpha/Delta variants circulating in the population increased from 1.7% in the week of 12 January to 72.6% in the week ending 12 June (Fig. 2C and movie S1). At the same time, the percentage of respondents living with school-aged children that reported having received any number of COVID-19 vaccine doses increased from 8.7 to 74.4% (Fig. 2D and movie S1).

Monthly rates of in-person schooling, vaccination, and average number of mitigation measures varied widely across counties, and this variation persisted throughout the spring semester (Fig. 2, A, B, and D).

The percent of respondents living with school-aged children that reported COVID-19–like illness (CLI; defined as fever of at least 37°C and cough, shortness of breath, or difficulty breathing) remained similar across the study period, ranging from 1.83% in January to 1.99% in June (table S2). The more specific indicators of potential SARS-CoV-2 infection, loss of taste and/or smell and a positive SARS-CoV-2 test with the past 14 days, decreased from 3.65 to 2.17% and 3.44 to 0.09%, respectively (table S2). These temporal trends were similar among all survey respondents and among respondents stratified by in-person schooling status, although respondents living with school-aged children not participating in any in-person schooling saw a decrease in CLI over the study period from 1.36 to 0.97% (table S2).

Overall—from 12 January to 12 June 2021—after adjusting for county-level SARS-CoV-2 biweekly attack rates averaged over the past 4 weeks, COVID-19 vaccination status, and other individual- and county-level factors, living in a household with a child in full-time in-person schooling was associated with increased odds of CLI [adjusted odds ratio (aOR), 1.32; 95% confidence interval (CI): 1.25 to 1.40], losing taste and/or smell (aOR, 1.19; 95% CI: 1.15 to 1.24), and reporting a positive SARS-CoV-2 test (aOR, 1.32; 95% CI: 1.27 to 1.38) (Fig. 3). In contrast to our previous analysis (9), we saw no clear trends by grade in any of the COVID-19–related outcomes associated with in-person schooling (Fig. 3 and fig. S2).

Consistent with our previous study (9), the most commonly reported mitigation measure was student masking, reported by 88% of respondents with in-person schooling in January and 84% in June (Fig. 4A). The second most common measure was teacher masking at 76% in January and 62% in June (Fig. 4A). Among those with in-person schooling, teacher masking was associated with the greatest risk reduction across all COVID-19–related outcomes, followed by

daily symptom screens, student masking, and restricted entry (Fig. 4B). These trends were consistent over time, with larger uncertainty in risk estimates in May to June (fig. S3). These patterns were also consistent for other households and individual-level COVID-19–related outcomes (fig. S4).

Across spring semester 2021, the association between CLI and loss of taste/smell and full-time and part-time in-person schooling disappeared with a report of four or more school-based mitigation measures, and risk of a SARS-CoV-2–positive test disappeared with a report of seven or more mitigation measures (Fig. 5). Results were consistent for other household- and individual-level COVID-19–related outcomes (fig. S5). These patterns were consistent from January to February and March to April (Fig. 5). By May to June, risks of all COVID-19–related outcomes disappeared when four or more mitigation measures were reported (Fig. 5).

Over the study period, each 10% increase in the state-level prevalence of Alpha/Delta variants was associated with increased baseline risk of CLI (aOR, 1.05; 95% CI: 1.03 to 1.07), loss of taste and/or smell (aOR, 1.02; 95% CI: 1.01 to 1.03), and reporting a positive SARS-CoV-2 test (aOR, 1.05; 95% CI: 1.03 to 1.06) after adjusting for background incidence, vaccination status, and other individual- and county-level characteristics. The rise of Alpha/Delta variants did not change the relative association between in-person schooling and COVID-19–related outcomes (Fig. 6). These findings did not change noticeably when we additionally adjusted for cumulative incidence of confirmed SARS-CoV-2 as an indirect indicator of population immunity (fig. S6).

We next examined whether and how the relationship between respondent vaccination status and living in a household with a child in in-person schooling varied over time (Fig. 7). We found that, by the end of the study period, there was a modest correlation between increased in-person schooling and lower vaccination rates at the county level (Pearson's $r = -0.15$; 95% CI: -0.18 to -0.12), while there was a positive correlation between vaccination rates and the number of

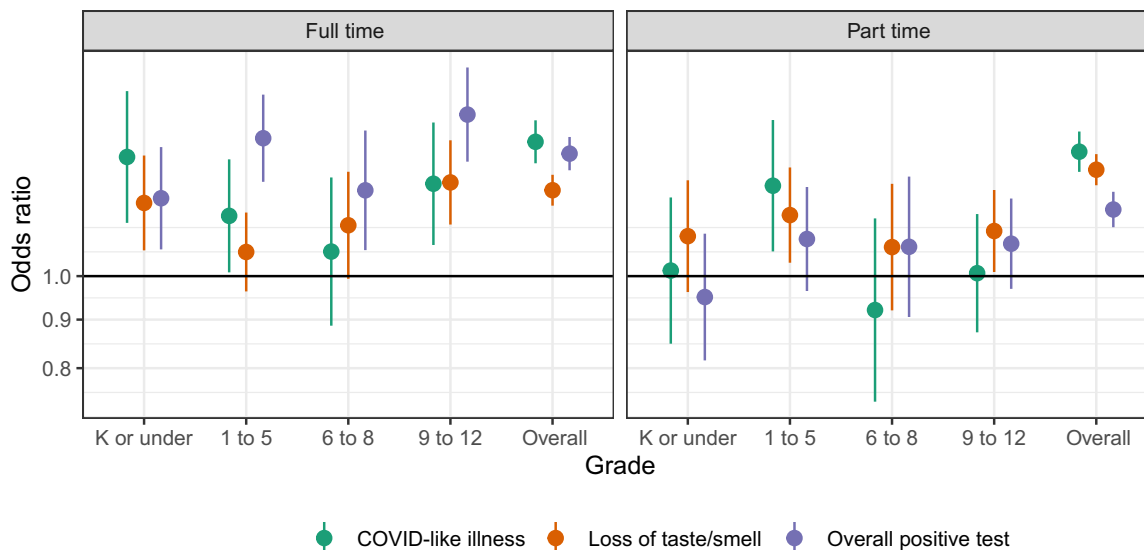


Fig. 3. Risk associated with in-person schooling overall and by grade level. Odds ratios of COVID-19–related outcomes by full- and part-time in-person schooling compared to no in-person schooling, overall and stratified by grade, adjusted for individual- and county-level covariates across the study period from 12 January to 12 June. Note that grade-stratified analyses include only those respondents who reported living with school-aged children in a single grade category, while the overall analyses include all respondents living with school-aged children; thus, the overall estimates do not always fall in between the grade-specific estimates.

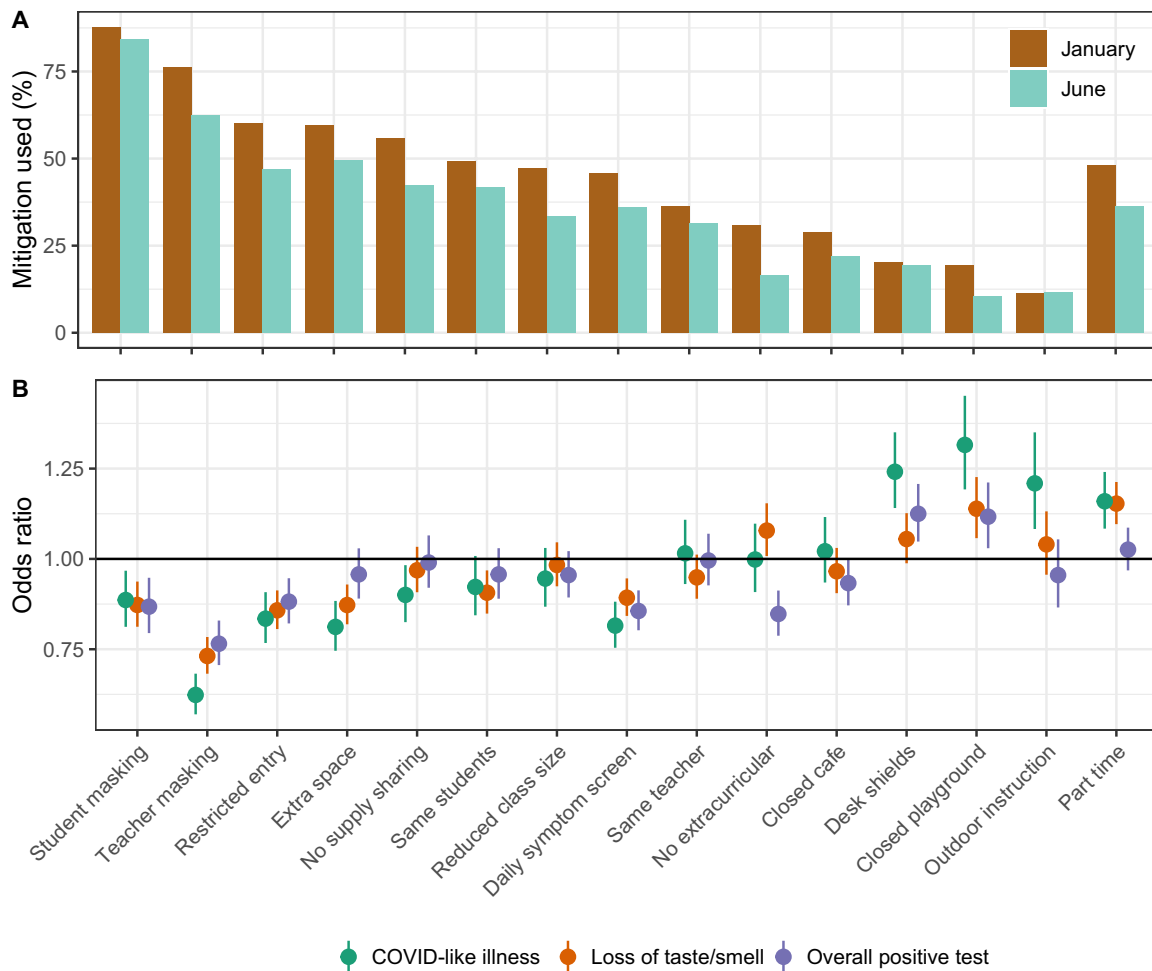


Fig. 4. Individual mitigation measures. (A) Percent of respondents with in-person schooling that reported each mitigation measure being used in January and June, weighted using U.S. CTIS survey weights. (B) Odds ratio of COVID-19–related outcomes among respondents with children in in-person schooling who reported each mitigation measure compared to those with children in in-person schooling who did not report each measure, adjusted for individual- and county-level covariates across the study period from 12 January to 12 June.

school-based mitigation measures reported in a county (Pearson’s $r = 0.36$; 95% CI: 0.32 to 0.39).

We explored the combined impact of living with a child in in-person schooling and vaccination status on COVID-19–related outcomes and whether interactions existed between the two risk factors (Fig. 8 and table S7). Overall, we found a reduction in the odds of COVID-19–related outcomes associated with both no in-person schooling and vaccination, with the latter having the far larger impact (Fig. 8A). Those having received two vaccine doses and not engaged in in-person schooling had the lowest risk of reporting COVID-19–related outcomes by a large margin (Fig. 8A). In-person schooling did not modify the association between vaccination status and reporting CLI but was, unexpectedly, associated with some variation in the apparent impact of vaccination on loss of taste/smell and reporting a positive SARS-CoV-2 test (Fig. 8B). The relative increase in the odds of reporting COVID-19–related outcomes associated with in-person schooling was the same regardless of vaccination status (Fig. 8C).

To understand COVID-19–related risk among teachers, we analyzed data from the 116,014 K-12 teachers included in the full U.S.

CTIS survey, whether or not they lived with a child participating in in-person schooling. We found that 86.0% of K-12 teachers reported work for pay conducted outside of their home in the previous 4 weeks. The percentage of teachers reporting paid work outside of the home increased from 77.5% in January to 92.4% in June, which was the largest increase in work outside the home for any occupation group (fig. S7). In comparison, the proportion of office and administrative support professionals that reported working outside of the home increased by a much smaller amount from 61.5 to 66.9% (fig. S7).

Overall, being a K-12 teacher conducting any work outside the home was associated with higher risk of losing taste/smell (aOR, 1.37; 95% CI: 1.13 to 1.65) and receiving a positive SARS-CoV-2 test (aOR, 2.04; 95% CI: 1.67 to 2.48) compared to K-12 teachers working exclusively from home (Fig. 9). However, we found no differences in the risk of reporting COVID-19–related outcomes between teachers and office and administrative support professionals working outside the home (other professions also had similar risk). These trends were consistent across time periods (fig. S8 and table S8). Notably, vaccinated teachers working outside the home were less

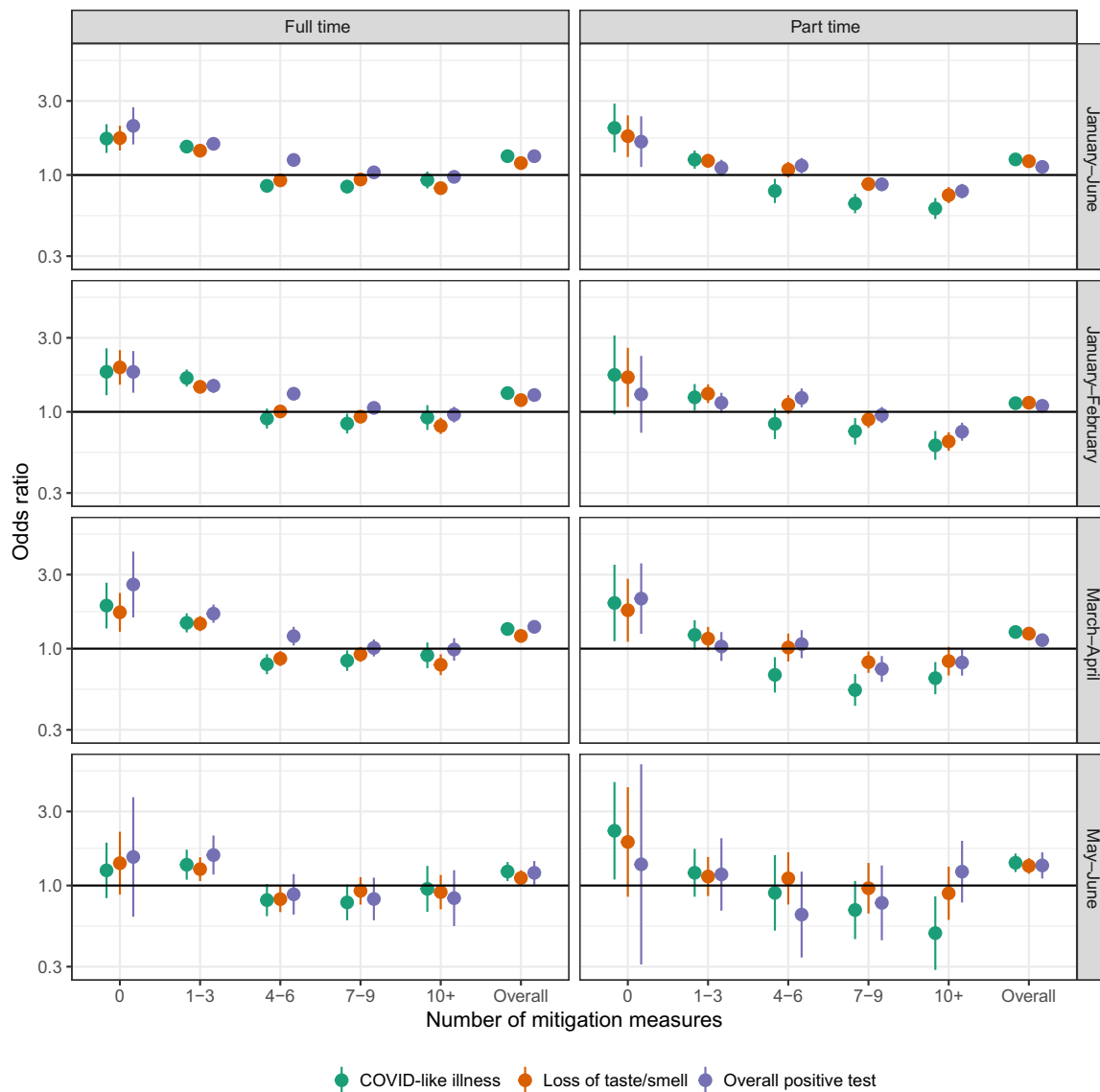


Fig. 5. Risk associated with in-person schooling over time by number of reported mitigation measures. Odds ratio of COVID-19–related outcomes by full- and part-time in-person schooling and number of school-based mitigation measures compared to no in-person schooling, adjusted for individual- and county-level covariates. “Overall” indicates adjusted odds ratios of full-time and part-time schooling with any number of mitigation measures compared to no in-person schooling. Odds ratios are shown across the entire study period from 12 January to 12 June (January–June) as well as within time periods from 12 January to 28 February (January–February), 1 March to 30 April (March–April), and 1 May to 12 June (May–June).

likely to report COVID-19–related outcomes than unvaccinated teachers reporting no work outside the home (fig. S9).

DISCUSSION

In-person schooling increased across the United States over the spring semester 2021, a period that also saw increasing COVID-19 vaccination rates and the spread of SARS-CoV-2 Alpha and Delta variants. In this study, we show that, despite these changes, associations seen in winter 2020–2021 (9) hold: in-person schooling is associated with increased risk of COVID-19–related outcomes in the household, but this risk can be reduced or eliminated with implementation of multiple mitigation measures in schools (Fig. 5 and

fig. S5). These findings were consistent across the study period even as vaccination rates increased, emphasizing the importance of layered mitigation measures to reduce the risk of transmission in schools. These measures remain crucial in light of increased COVID-19 cases due to the Delta and, now, Omicron variants (3) as well as the potential for future outbreaks and are consistent with CDC guidance (16).

We also found that changes in in-person schooling and mitigation measures varied between states and counties, which has likely contributed to heterogeneity in risk. We found that higher levels of in-person schooling were accompanied by lower levels of mitigation measures (Figs. 1, 2, and 7). We also found that, from March to April, and even more so from May to June, vaccination rates were

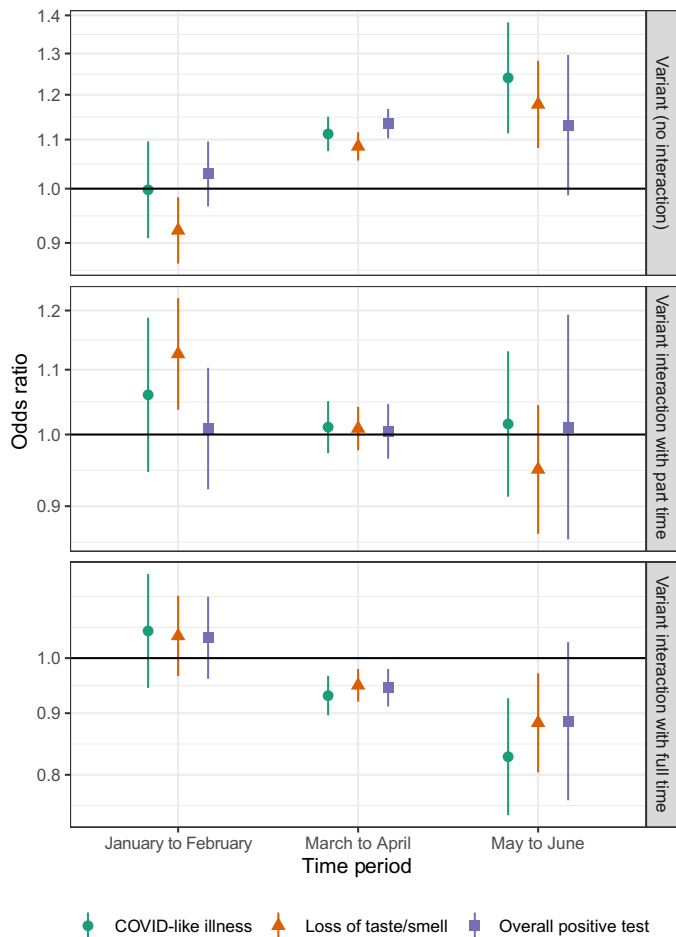


Fig. 6. Risk associated with in-person schooling and Alpha/Delta variants over time. Estimated baseline risk (top) associated with the proportion of cases due to the Alpha and Delta variants and interaction with part-time (middle) and full-time (bottom) in-person schooling status. Interaction terms show the additional impact of the variant on top of baseline risk due to part- and full-time in-person schooling. Risk is shown within a 2-month time strata.

positively correlated with mitigation measures and inversely correlated with in-person schooling (Fig. 7). This tendency of communities to eschew both kinds of control may lead to significant increases in overall COVID-19 risk in these areas.

We found that being unvaccinated and living with a student engaged in in-person schooling was associated with the highest risk of reporting COVID-19-related outcomes, while, expectedly, those reporting two doses of vaccine and no one in the household engaged in in-person schooling had the lowest risk (Fig. 8A). Even among individuals with two vaccine doses, in-person schooling was associated with significantly increased risk of COVID-19-related outcomes in household members compared to individuals with no children in in-person schooling (Fig. 8C). In other words, our results suggest that, although adult vaccination substantially reduces the overall risk from living with a child engaged in in-person schooling, the relative change in risk due to in-person schooling is similar to that seen in the unvaccinated.

The proportion of teachers working outside the home increased more than any other professional group over the spring semester 2021 (fig. S7). Consistent with the in-person schooling results, K-12 teachers working for pay outside the home were at increased risk of COVID-19-related outcomes (Fig. 9), although the additional risk was similar to that in other occupations (Fig. 9). Vaccinated K-12 teachers working outside the home reported fewer COVID-19-related outcomes than the unvaccinated not working outside their homes. This emphasizes the critical role that vaccination can play in a safe return to the classroom for teachers.

Furthermore, similar to previous work (9), we found that teacher masking, student masking, restricted entry, and symptom screens were individually associated with the greatest reduction in the risk associated with in-person schooling (Fig. 4). We also found a particularly strong signal for CLI for teacher masking, which may indicate that masks are preventing the spread of other respiratory infections. In addition, we found that some of the less commonly used mitigation measures, such as desk shields and closed playgrounds, were associated with increased risk (Fig. 4). This may reflect reduced utility of these measures and/or a saturation effect, since these are often in place alongside other mitigation approaches. Closed extracurricular

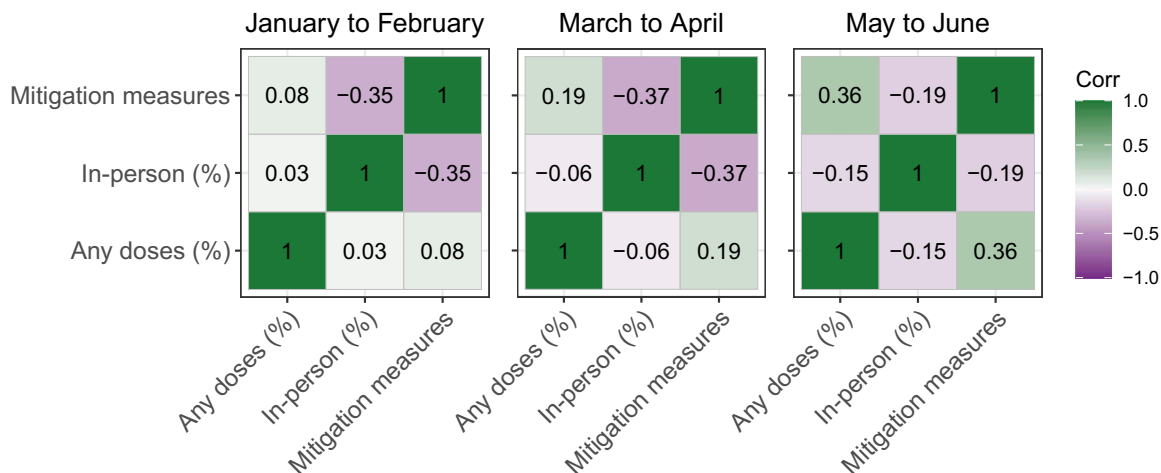


Fig. 7. Correlations among vaccination, in-person schooling, and mitigation measures. Pearson's correlation coefficients between percent vaccinated with at least one dose of a COVID-19 vaccine, percent of respondents living with school-aged children reporting in-person schooling, and average number of school-based mitigation measures by county and month. Results are shown within time periods of 12 January to 28 February (left), 1 March to 30 April (center), and 1 May 1 to 12 June (right). All coefficients shown were significant ($P < 0.05$). Limited to county-months with at least five respondents.

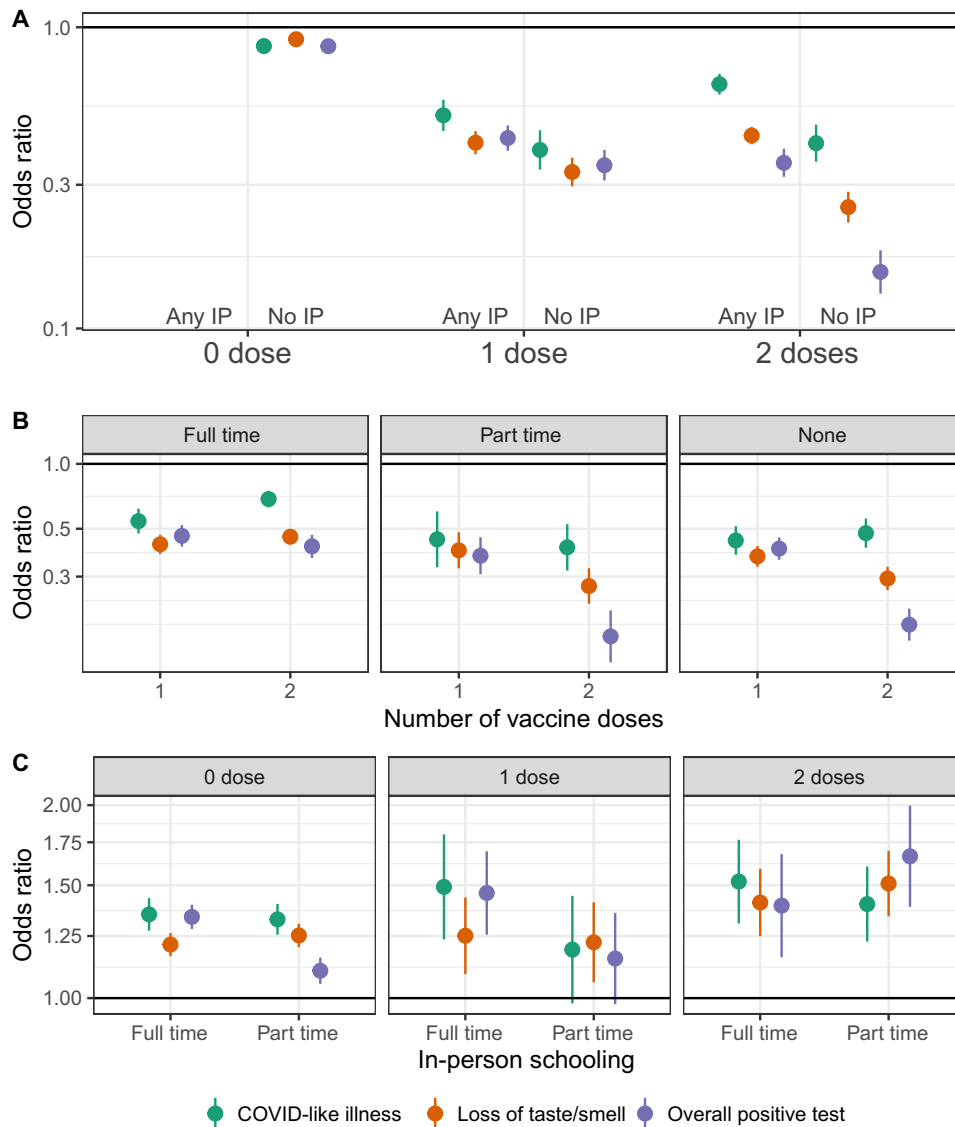


Fig. 8. Risk associated with in-person schooling and number of vaccine doses. (A) Odds ratio of COVID-19–related outcomes by number of reported vaccine doses (0, 1, or 2) and in-person schooling status [living in a household with a child participating in any in-person (any IP) or no in-person (no IP) schooling], adjusted for individual- and county-level covariates. Zero vaccine doses and any in-person schooling is the reference group. (B) Odds ratio of COVID-19–related outcomes with one and two vaccine doses compared to zero doses when data are stratified by none, part-time, and full-time in-person schooling. (C) Odds ratio of COVID-19–related outcomes with part- and full-time in-person schooling compared to no in-person schooling when data are stratified by vaccine doses.

activities were associated with reduced risk of positive tests in both respondents and family members but not CLI or loss of taste/smell (Fig. 4 and fig. S4), which could reflect increased testing in households with students in extracurricular activities.

Although our results do not indicate that the rise of the Alpha variant over the study period changed the risk associated with in-person schooling, there were some unexpected patterns in the interaction between full-time schooling and variant prevalence, which may reflect a limitation of the data. As the prevalence of the Alpha variant was increasing in all states over the study period, as well as within each time strata that we investigated, some of this effect may be due to other time-varying factors that are not captured in the data such as changing compliance with intervention measures. In

addition, data on variant prevalence were only available at the state level, which may obscure important county-level differences.

We were also unable to evaluate patterns associated with the rise of Delta and Omicron variants, which have dominated the 2021–2022 school year (3). However, as Delta and Omicron are more transmissible than Alpha in vaccinated and unvaccinated populations (17, 18), layered mitigation measures are likely even more important for reducing the risk now posed by in-person schooling. This is consistent with results from a model-based simulation study in the California Bay Area, which estimated that both vaccination and masking would be required to reduce, but not eliminate, Delta infections in students in fall semester 2021 and that additional mitigation measures would be required to further reduce risk (19). Investigating the impact of

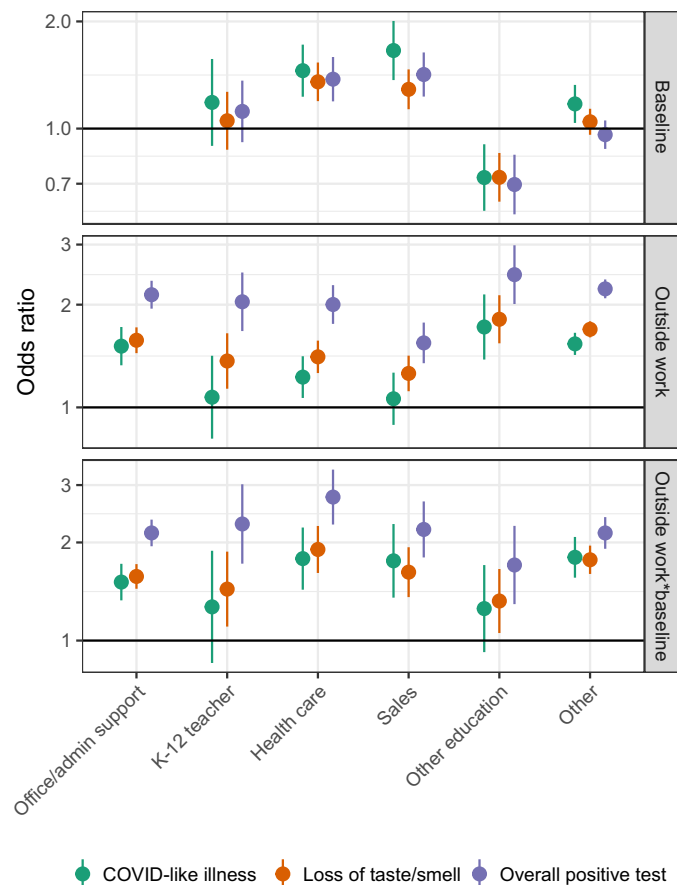


Fig. 9. Risk by occupation and paid work outside the home. Odds ratio of COVID-19–related outcomes contrasting office workers not reporting extra household work for pay to those in other employment categories not reporting work for pay outside the home (top) and to those reporting work for pay outside the home (bottom). The middle panel shows the odds ratio (i.e., increased risk) within each category associated with working outside the home compared to no work outside the home. Food service workers are excluded from this analysis, because less than 5% reported working exclusively from home (fig. S7).

variants on in-person schooling risk for the 2021–2022 school year will be an important area for future work.

This study has several additional limitations. The survey data that informed the risk estimates were self-reported and subject to recall bias. They were also gathered through the Facebook platform and may not be representative of the underlying populations, although this should be accounted for at least in part through the survey weights (20). We previously investigated this further and found that in-person schooling estimated via U.S. CTIS was highly correlated with county-level rates reported for public schools (21). Here, we found that vaccination rates were also highly correlated with CDC’s reported data, although the survey data consistently reported higher levels of vaccination than the CDC (fig. S1) (15), which has been described by others (22). While CDC’s reported vaccination data are known to be underreported, this could also reflect a bias in the survey data (22). For example, U.S. CTIS may be capturing a more affluent, less rural population, which could explain, in part, why we found particularly large differences in the survey data compared to the CDC data in states with large rural populations (fig. S1B). Despite

this, our current results were consistent across a range of COVID-19–related outcomes (figs. S4 and S5), and our previous results were consistent after adjusting for geographic- and individual-based propensity scores as well as within strata of reported schooling behaviors, states, percent white, poverty, and access to broadband internet (9). Moreover, any biases in survey representativeness would have to be differential by schooling status to alter our conclusions.

We were also limited by data available to us in the survey or at the county level. In our analysis of risk by occupation, we were unable to examine risk by the amount of time spent working outside of the home. In addition, there may be confounding factors that we were not able to adjust for, such as community-level vaccination and immunity. Although, in addition, adjusting for cumulative incidence as a proxy for this did not noticeably affect our results (fig. S6). We were also not able to evaluate the timing of vaccine doses in relation to COVID-19–related outcomes. Thus, it is possible that some of the individuals who received two vaccine doses had been infected before developing immunity. Last, of the seven COVID-19–related outcomes that we measured, none specifically assessed asymptomatic infection. In addition, SARS-CoV-2 test positivity requires seeking out a test, and CLI is not specific for COVID-19. That our findings were largely concordant between these varied outcomes and supports our overall conclusions. Some of the differences we did find between CLI and SARS-CoV-2–positive tests could reflect increases in non-SARS-CoV-2 respiratory infections over the spring semester, as was found in a study in Hong Kong (23).

Overall, our findings support previous studies that have shown secondary transmission and outbreaks associated with in-person schooling and child care (24–28). While there are other studies that have shown relatively low risk of transmission in schools (29–33), SARS-CoV-2 mitigation measures were in place in each of these settings. Thus, there is abundant evidence to indicate that in-person schooling plays a role in SARS-CoV-2 transmission, but this risk can be mitigated. Despite major changes in in-person schooling behavior and the epidemiological situation over the course of the spring semester 2021, the apparent relative risks associated with in person schooling—and the measures that worked to mitigate these risks—remained mostly the same. Hence, as we confront the current and possible future COVID-19 epidemics, these results can help guide us as we strive to keep our homes safe while not sacrificing the education of our children.

MATERIALS AND METHODS

Survey data

We analyzed survey data collected by the U.S. CTIS. These are openly available human data, which can be obtained at <https://cmu-delphi.github.io/delphi-epidata/>. The U.S. CTIS is administered by the Delphi Group in partnership with Facebook and includes questions on demographics, symptoms related to COVID-19, positive SARS-CoV-2 test results, vaccination, and schooling for any children in the household (7); detailed survey questionnaires are available in (8). The U.S. CTIS uses a two-stage sampling (20, 34) and provides survey weights to allow adjustment for (i) differences between the U.S. population and U.S. Facebook users and (ii) the propensity of a Facebook user to take the survey. Previously, we found that alternative weighting and stratification of the U.S. CTIS survey data did not have a qualitative impact on estimated risk posed by in-person schooling (9); therefore, we used the U.S. CTIS–provided survey

weights to adjust the data throughout this study. Our study period encompasses waves 7 and 8 and 10 and 11 of this survey, covering the spring semester of the 2020–2021 school year from 1 January to 12 June 2021. Additional details can be found in Supplementary Methods.

Variant data

Daily state-level SARS-CoV-2 variant data were obtained from outbreak.info (35), which uses data from Global Initiative on Sharing All Influenza Data (GISAID) (13). These data include the daily proportion of sequences for each variant. For each state, a smooth spline with six degrees of freedom was fitted to the combined Alpha and Delta variant proportion over time. Each survey response was assigned the fitted variant proportion for their state corresponding to 7 days before the survey response date.

CDC vaccine data

Daily county-level vaccination data were obtained from the CDC (15) on 14 July 2021. Weekly proportions of individuals in each county that had received at least one vaccine dose were calculated by taking into account county populations and percent completeness of the reported CDC data. These data were used in linear regression and Pearson's correlation analyses to examine how county-level vaccination rates reported by the CDC compared with those reported by the Facebook survey respondents living in a household with school-aged children. The latter survey data were used as individual-level covariates to adjust for respondent vaccination status in the analyses described below and in Supplementary Methods.

COVID-19 case data

County-level COVID-19 case data were obtained through the Johns Hopkins Center for Systems Science and Engineering COVID-19 Dashboard (36). County-week-level attack rates were calculated as the average 2-week incidence in the past 4 weeks. Attack rates were calculated by dividing reported COVID-19 cases by 2020 county populations, which were obtained from the U.S. Census Bureau using the tidycensus package (<https://cran.r-project.org/package=tidycensus>). Adjustment for attack rates was done on the basis of the log base 2 cases per thousand [i.e., $\log_2(\{\text{average biweekly attack rate per 1000}\} + 1)$].

Covariates and outcomes

In addition to average 2-week incidence in the past 4 weeks, county-level covariates (obtained from the 2014–2018 American Community Survey using the tidycensus package) included the total population, percent of the population that was white, percent of households with income below the poverty threshold, a measure of income inequality, and metropolitan type. Individual- and household-level covariates (obtained from the U.S. CTIS dataset) included gender, age, occupation, educational level, household size, masking behavior, out-of-state travel, vaccination with any number of COVID-19 vaccine doses, whether they reported a visit to a bar/restaurant/cafe to an event with more than 10 people, and whether they used public transit. Primary outcomes included CLI, loss of taste and/or smell, and a positive SARS-CoV-2 test results in the past 24 hours. Secondary outcomes included CLI in any household member, contact with a household member who received a positive test result, and a positive test result when the test was not indicated. Detailed descriptions of all variables and outcomes can be found in Supplementary Methods.

Analysis

All analyses were conducted using quasi-binomial regression accounting for survey weights using the *srvyr* package (<https://cran.r-project.org/package=srvyr>) in the R statistical language. The overall analysis included adjustments for the baseline covariates described above as well as full- and part-time in-person schooling status of children living in the household. The analysis of COVID-19 risk and number of school-based mitigation measures included full- and part-time in-person schooling variables categorized by number of mitigation measures. The analysis of individual control measures was restricted to respondents living with children in in-person schooling and adjusted for baseline covariates as well as each individual school-based mitigation measure. The analysis of in-person schooling and Alpha/Delta variants included the baseline covariates, the fitted combined Alpha/Delta variant proportion, and two interaction terms between variant proportion and each of full- and part-time in-person schooling. The analyses of in-person schooling and number of vaccine doses included the baseline covariates but with vaccination status replaced by the number of COVID-19 vaccine doses. The analysis of the risk of COVID-19-related outcomes among educational professionals was conducted among all respondents, regardless of living with a child in in-person schooling, and included baseline covariates as well as an interaction term between occupation types and an indicator for any paid work outside the home in the past 4 weeks. Full details of each analysis including regression formulas can be found in Supplementary Methods.

SUPPLEMENTARY MATERIALS

Supplementary material for this article is available at <https://science.org/doi/10.1126/sciadv.abm9128>

REFERENCES AND NOTES

1. "Map: Where were schools required to be open for the 2020-21 school year?," *Education Week*, 28 July 2020; www.edweek.org/leadership/map-where-are-schools-closed/2020/07.
2. WHO, WHO coronavirus (COVID-19) dashboard; access date: 11 September 2021, <https://covid19.who.int>.
3. Centers for Disease Control and Prevention (CDC), COVID data tracker (CDC, 2020); <https://covid.cdc.gov/covid-data-tracker>.
4. MCH Strategic Data, COVID-19 IMPACT: School district operational status; access date: 11 September 2021, <https://mchstrategicdata.maps.arcgis.com/apps/MapSeries/index.html?appid=f49763c1f2c84d8bbb78e3e6b5ed5477>.
5. R. M. Viner, O. T. Mytton, C. Bonell, G. J. Melendez-Torres, J. Ward, L. Hudson, C. Waddington, J. Thomas, S. Russell, F. van der Klis, A. Koirala, S. Ladhani, J. Panovska-Griffiths, N. G. Davies, R. Booy, R. M. Eggo, Susceptibility to SARS-CoV-2 infection among children and adolescents compared with adults: A systematic review and meta-analysis. *JAMA Pediatr.* **175**, 143–156 (2021).
6. Delphi Epidata API, COVID-19 trends and impact Survey (Delphi Research Group at Carnegie Mellon University, 2020); <https://cmu-delphi.github.io/delphi-epidata/api/covidcast-signals/fb-survey.html>.
7. F. Kreuter, N. Barkay, A. Bilinski, A. Bradford, S. Chiu, R. Elliot, J. Fan, T. Gallii, D. Haimovich, B. Kim, S. LaRocca, Y. Li, K. Morris, S. Presser, T. Sarig, J. A. Salomon, K. Stewart, E. A. Stuart, R. Tibshirani, Partnering with a global platform to inform research and public policy making. *Survey Research Methods.* **14**, 159–163 (2020).
8. Delphi Epidata API, COVIDcast Epidata API (Delphi Research Group at Carnegie Mellon University); access date: 10 September 2021, <https://cmu-delphi.github.io/delphi-epidata/api/covidcast.html>.
9. J. Lessler, M. K. Grabowski, K. H. Grantz, E. Badillo-Goicoechea, C. J. E. Metcalf, C. Lupton-Smith, A. S. Azman, E. A. Stuart, Household COVID-19 risk and in-person schooling. *Science* **372**, 1092–1097 (2021).
10. S. M. Moghadas, T. N. Vilches, K. Zhang, C. R. Wells, A. Shoukat, B. H. Singer, L. A. Meyers, K. M. Neuzil, J. M. Langley, M. C. Fitzpatrick, A. P. Galvani, The impact of vaccination on coronavirus disease 2019 (COVID-19) outbreaks in the United States. *Clin. Infect. Dis.* **73**, 2257–2264 (2021).

11. E. Mathieu, H. Ritchie, E. Ortiz-Ospina, M. Roser, J. Hasell, C. Appel, C. Giattino, L. R.-Guirao, A global database of COVID-19 vaccinations. *Nat. Hum. Behav.* **5**, 947–953 (2021).
12. N. G. Davies, S. Abbott, R. C. Barnard, C. I. Jarvis, A. J. Kucharski, J. D. Munday, C. A. B. Pearson, T. W. Russell, D. C. Tully, A. D. Washburne, T. Wenseleers, A. Gimma, W. Waites, K. L. M. Wong, K. van Zandvoort, J. D. Silverman; CMMID COVID-19 Working Group; COVID-19 Genomics UK (COG-UK) Consortium, K. D.-Ordaz, R. Keogh, R. M. Eggo, S. Funk, M. Jit, K. E. Atkins, W. J. Edmunds, Estimated transmissibility and impact of SARS-CoV-2 lineage B.1.1.7 in England. *Science* **372**, eabg3055 (2021).
13. GISAI, hCov19 variants; access date: 15 July 2021, www.gisaid.org/hcov19-variants.
14. H. Allen, A. Vusirikala, J. Flannagan, K. A. Twohig, A. Zaidi, N. Groves, J. Lopez-Bernal, R. Harris, A. Charlett, G. Dabrera, M. Kall, Increased household transmission of COVID-19 cases associated with SARS-CoV-2 variant of concern B.1.617.2: A national case-control study. *Lancet Reg Health Eur.* **12**, 100252 (2021).
15. Centers for Disease Control and Prevention (CDC), COVID-19 vaccinations in the United States, county (CDC, 2022); <https://data.cdc.gov/Vaccinations/COVID-19-Vaccinations-in-the-United-States-County/8xx-amqh>.
16. Centers for Disease Control and Prevention (CDC), Guidance for COVID-19 prevention in K-12 schools and ECE programs (CDC, 2021); www.cdc.gov/coronavirus/2019-ncov/community/schools-childcare/k-12-guidance.html.
17. Centers for Disease Control and Prevention (CDC), Delta variant: What we know about the science (CDC, 2020); www.cdc.gov/coronavirus/2019-ncov/variants/delta-variant.html.
18. N. Andrews, J. Stowe, F. Kirsebom, S. Toffa, T. Rickeard, E. Gallagher, C. Gower, M. Kall, N. Groves, A.-M. O'Connell, D. Simons, P. B. Blomquist, A. Zaidi, S. Nash, N. I. B. A. Aziz, S. Thelwall, G. Dabrera, R. Myers, G. Amirthalingam, S. Gharbia, J. C. Barrett, R. Elson, S. N. Ladhani, N. Ferguson, M. Zambon, C. N. Campbell, K. Brown, S. Hopkins, M. Chand, M. Ramsay, J. L. Bernal, Effectiveness of COVID-19 vaccines against the Omicron (B.1.1.529) variant of concern. medRxiv 21267615 (2021).
19. J. R. Head, K. L. Andrejko, J. V. Remais, Model-based assessment of SARS-CoV-2 Delta variant transmission dynamics within partially vaccinated K-12 school populations. *Lancet Reg. Health Am.* **2021**, 100133 (2021).
20. N. Barkay, C. Cobb, R. Eilat, T. Gallii, D. Haimovich, S. LaRocca, K. Morris, T. Sarig, Weights and methodology brief for the COVID-19 symptom survey by University of Maryland and Carnegie Mellon University, in partnership with Facebook. arXiv:14675 (2020).
21. C. Lupton-Smith, E. B. Goicoechea, M. Collins, J. Lessler, M. K. Grabowski, E. A. Stuart, Consistency between household and county measures of K-12 onsite schooling during the COVID-19 pandemic. arXiv:13296 (2021).
22. V. C. Bradley, S. Kuriwaki, M. Isakov, D. Sejdinovic, X.-L. Meng, S. Flaxman, Unrepresentative big surveys significantly overestimated US vaccine uptake. *Nature* **600**, 695–700 (2021).
23. M. W. Fong, N. H. L. Leung, B. J. Cowling, P. Wu, Upper respiratory infections in schools and childcare centers reopening after COVID-19 dismissals, Hong Kong. *Emerg. Infect. Dis.* **27**, 1525–1527 (2021).
24. J. A. W. Gold, Clusters of SARS-CoV-2 infection among elementary school educators and students in one school district—Georgia, December 2020–January 2021. *MMWR Morb. Mortal. Wkly Rep.* **70**, 289–292 (2021).
25. C. M. Szablewski, K. T. Chang, M. M. Brown, V. T. Chu, A. R. Yousaf, N. Anyalechi, P. A. Aryee, H. L. Kirking, M. Lumsden, E. Mayweather, C. J. McDaniel, R. Montierth, A. Mohammed, N. G. Schwartz, J. A. Shah, J. E. Tate, E. Dirlikov, C. Drenzek, T. M. Lanzieri, R. J. Stewart, SARS-CoV-2 transmission and infection among attendees of an overnight camp—Georgia, June 2020. *MMWR Morb. Mortal. Wkly Rep.* **69**, 1023–1025 (2020).
26. A. S. Lopez, M. Hill, J. Antezano, D. Vilven, T. Rutner, L. Bogdanow, C. Clafin, I. T. Kracalik, V. L. Fields, A. Dunn, J. E. Tate, H. L. Kirking, T. Kiphibane, I. Risk, C. H. Tran, Transmission dynamics of COVID-19 outbreaks associated with child care facilities—Salt Lake City, Utah, April–July 2020. *MMWR Morb. Mortal. Wkly Rep.* **69**, 1319–1323 (2020).
27. C. S.-Zamir, N. Abramson, H. Shooab, E. Libal, M. Bitan, T. Cardash, R. Cayam, I. Miskin, A large COVID-19 outbreak in a high school 10 days after schools' reopening, Israel, May 2020. *Euro Surveill.* **25**, 2001352 (2020).
28. T. Lam-Hine, Outbreak associated with SARS-CoV-2 B.1.617.2 (Delta) variant in an elementary school—Marin County, California, May–June 2021. *MMWR Morb. Mortal. Wkly Rep.* **70**, 1214–1219 (2021).
29. S. A. Ismail, V. Saliba, J. Lopez Bernal, M. E. Ramsay, S. N. Ladhani, SARS-CoV-2 infection and transmission in educational settings: A prospective, cross-sectional analysis of infection clusters and outbreaks in England. *Lancet Infect. Dis.* **21**, 344–353 (2021).
30. K. O. Zimmerman, I. C. Akinboyo, M. A. Brookhart, A. E. Boutzoukas, K. A. McGann, M. J. Smith, G. M. Panayotti, S. C. Armstrong, H. Bristow, D. Parker, S. Zadrozny, D. J. Weber, D. K. Benjamin, For the Abc science collaborative, incidence and secondary transmission of SARS-CoV-2 infections in schools. *Pediatrics* **147**, 2021 (2021).
31. A. Falk, COVID-19 cases and transmission in 17 K-12 Schools—Wood County, Wisconsin, August 31–November 29, 2020. *MMWR Morb. Mortal. Wkly Rep.* **70**, 136–140 (2021).
32. R. B. Hershov, Low SARS-CoV-2 transmission in elementary schools—Salt Lake County, Utah, December 3, 2020–January 31, 2021. *MMWR Morb. Mortal. Wkly Rep.* **70**, 442–448 (2021).
33. J. K. Varma, J. Thamkittikasem, K. Whittemore, M. Alexander, D. H. Stephens, K. Arslanian, J. Bray, T. G. Long, COVID-19 infections among students and staff in New York City public schools. *Pediatrics* **147**, e2021050605 (2021).
34. Delphi Epidata API, Questions and coding (Delphi Research Group at Carnegie Mellon University, 2021); <https://cmu-delphi.github.io/delphi-epidata/symptom-survey/coding.html>.
35. outbreak.info; access date: 28 July 2021, <https://outbreak.info>.
36. E. Dong, H. Du, L. Gardner, An interactive web-based dashboard to track COVID-19 in real time. *Lancet Infect. Dis.* **20**, 533–534 (2020).

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Data and materials availability: All data needed to evaluate the conclusions in the paper are present in the paper and/or the Supplementary Materials. Survey data are freely available from the CMU Delphi Research Group to researchers at universities and nonprofits at <https://cmu-delphi.github.io/delphi-epidata/>. All analytic code with dummy datasets is available at <https://doi.org/10.5281/zenodo.5942681> and <https://github.com/HopkinsIDD/inperson-schooling-covid-spring-2021>.

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