Rejection-free quantum Monte Carlo in continuous time from transition path sampling

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Continuous-time quantum Monte Carlo refers to a class of algorithms designed to sample the thermal distribution of a quantum Hamiltonian through exact expansions of the Boltzmann exponential in terms of stochastic trajectories which are periodic in imaginary time. Here, we show that for (sign-problem-free) quantum many-body systems with discrete degrees of freedom—such as spins on a lattice—this sampling can be done in a rejection-free manner using transition path sampling (TPS). The key idea is to converge the trajectory ensemble through updates where one individual degree of freedom is modified across all time while the remaining unaltered ones provide a time-dependent background. The ensuing single-body dynamics provides a way to generate trajectory updates exactly, allowing one to obtain the target ensemble efficiently via rejection-free TPS. We demonstrate our method on the transverse field Ising model in one and two dimensions, and on the quantum triangular plaquette (or Newman-Moore) model. We show that despite large autocorrelation times, our method is able to efficiently recover the respective quantum phase transition of each model. We also discuss the connection to rare event sampling in continuous-time Markov dynamics.

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I. INTRODUCTION

Statistical mechanics has facilitated the study of manybody systems through Monte Carlo sampling for several decades. Originally developed for classical Hamiltonians, classical Monte Carlo methods allow the sampling of the Boltzmann distribution by proposing updates to a configuration together with a criterion, such as Metropolis [1,2] or Glauber [3], to accept or reject the changes. At the same time, Monte Carlo sampling can be generalized to study the Boltzmann distribution of quantum many-body Hamiltonians [4]. However, a complication here is that the weights of the configurations require the evolution to be computed in imaginary time.

A standard approach is to split the imaginary-time evolution of the quantum partition sum into discrete time steps, giving rise to trajectories of configurations in discretized imaginary time, where the weightings of the trajectories can be estimated for small time steps through a Trotter-Suzuki decomposition [5,6]. This procedure can be made exact by considering a *continuous-time* expansion instead, in terms of a perturbative expansion of the Boltzmann exponential in the interaction picture and to all orders, yielding a sum over classical trajectories in continuous time [7–10]. In either of these approaches, there is added complexity in relation to the classical case, as imaginary-time trajectories have to

be sampled by making changes to the temporal degrees of freedom in addition to the spatial ones [11].

In this paper, we focus on the continuous-time expansion of the quantum Boltzmann distribution through trajectory ensembles. For systems with discrete local degrees of freedom (and in the absence of a sign problem [12,13]), we show how the trajectories that define the quantum partition sum can be sampled through the realization of the so-called *Doob dynamics* (see, e.g., Refs. [14–16]), which optimally samples rare events. Specifically, we show how this can be achieved efficiently through exact trajectory updates which are *local in space but extensive in time*. In this way we can generate ensembles of trajectories using a version of transition path sampling (TPS), which is *rejection-free* (in the sense it does not require any acceptance criteria) [17,18].

We illustrate the method by studying the transverse field Ising model (TFIM) in one and two spatial dimensions. In this case, we sample many-body trajectories by updating the trajectories of individual spins, while keeping the other spins as an effective time-dependent background. This approach is similar to the one in Ref. [17], which studied the TFIM on the Bethe lattice, and to the one in Ref. [18], which studied transition events of the classical 2D Ising model. It also generalizes the approach adopted in Ref. [19], which studied the groundstate properties of the 1D TFIM via time- and space-local TPS updates. Here, we demonstrate that this approach is capable of predicting the well-known continuous phase transitions for the 1D and 2D TFIMs, despite the fact that the local updates suffer from large autocorrelation times close to criticality. It is important to note that there exist update schemes which can overcome the difficulties of criticality, for example, cluster updates [20,21] (a generalization of the Swendsen-Wang algorithm

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[22] to continuous-time ensembles), and the Worm algorithm [23]. Reference [24] demonstrates the use of our approach where these alternative methods are difficult to formulate.

We also show how to use our approach with more complex spin interactions by means of customized local updates. As a concrete example, we study the quantum Newman-Moore model [19,24–29], also known as the quantum triangular plaquette model (QTPM). This model builds on the classical TPM [30–32], a system of spins with three-body interactions studied in the context of glasses, making it quantum by adding a transverse magnetic field. For the QTPM, we devise local updates by simultaneously changing the trajectories of three neighboring spins. Our method can effectively recover the first-order quantum phase transition of the QTPM [19,24– 26,28]. We further show how thermal annealing can be implemented to improve sampling close to the transition point.

The paper is structured as follows. Section II formalizes the connection between the quantum Boltzmann distribution and rare-trajectory sampling, explaining how observables can be estimated from trajectories. Section III then explains how the optimal Doob dynamics can be used to sample trajectories from the quantum Boltzmann distribution. Section IV demonstrates how this approach can be used to incorporate a single-spin update scheme for spin models with a simple transverse field, in analogy to Ref. [17], by implementing the method for the 1D and 2D TFIM. Section V generalizes the method to Hamiltonians with more complex interaction terms, using the QTPM as a concrete example. We give our conclusions in Sec. VI, where we discuss the possibility of more involved schemes, including updates which are collective in space. Appendix B demonstrates how to use our approach for rare event sampling in classical continuous-time Markov dynamics, where we determine the dynamical large deviations (LDs) of the TFIM and QTPM as an example.

II. MONTE CARLO SAMPLING OF THE BOLTZMANN DISTRIBUTION

A. The quantum partition function and continuous-time stochastic dynamics

We consider a system with a Hamiltonian \hat{H} and a discrete set of configurations that defines a basis $\{|x\rangle\}$ of its Hilbert space, $\mathcal{H}[33]$. In this basis, the Hamiltonian can be decomposed into a diagonal and an off-diagonal part, $\hat{H} = \hat{H}^c - \hat{K}$,

$$\hat{H}^{c} = \sum_{x} H_{x}^{c} |x\rangle \langle x|, \qquad (1)$$

$$\hat{K} = \sum_{x,y \neq x} \hat{K}_{x \to y} \equiv \sum_{x,y \neq x} K_{x \to y} |y\rangle \langle x|.$$
(2)

We will refer to operators which act on the whole Hilbert space by symbols with a hat (e.g., \hat{H}^c), and their individual matrix elements by symbols without a hat.

The statistical properties of \hat{H} at some finite temperature are characterized by the partition function:

$$Z_{\beta} = \operatorname{Tr}[e^{-\beta\hat{H}}] = \sum_{x} \langle x | e^{-\beta\hat{H}} | x \rangle.$$
(3)

Equation (3) can be expressed as a sum over stochastic paths by considering its Dyson series expansion [34],

$$Z_{\beta} = \sum_{M=0}^{\infty} \sum_{x_{0}, \cdots, x_{M}} \int dt_{M} \dots dt_{1} \langle x_{0} | e^{-(\beta - t_{M})\hat{H}^{c}} \\ \times \prod_{m=1}^{M} [\hat{K}_{x_{m-1} \to x_{m}} e^{-(t_{m} - t_{m-1})\hat{H}^{c}}] | x_{0} \rangle, \qquad (4)$$

where the integrals over times t_m are performed with the limits $\beta \ge t_M \ge \cdots \ge t_0$. Note that the product and integrals in Eq. (4) are omitted for the case of M = 0.

Each path in Eq. (4) can be interpreted as a classical stochastic trajectory in continuous time with time extent β , where the operator $\hat{K}_{x_{m-1} \to x_m}$ makes the instantaneous transition $x_{m-1} \to x_m$ at the time t_m . We denote a stochastic trajectory of time extent β with M jumps as

$$\mathbf{x} = \{(t_0, x_0), \dots, (t_M, x_M)\}.$$
 (5)

Note that the last jump occurs at $t_M \leq \beta$. We then represent Eq. (4) as a sum over these jumps,

$$Z_{\beta} = \sum_{\{x\}} \delta[x(0), x(\beta)] e^{-\int_{0}^{\beta} dt \, H_{x(t)}^{c}} \prod_{m=1}^{K(x)} K_{x_{m-1} \to x_{m}}, \quad (6)$$

where x(t) denotes the configuration of the trajectory x at time t, the delta function $\delta[x(0), x(\beta)]$ allows only for trajectories which are periodic in time, and $\mathcal{K}(x)$ is the *trajectory observable* (which will henceforth be denoted by a calligraphic font) which returns the number of transitions, M, which occur in the trajectory x. The sum over $\{x\}$ indicates a sum over all paths for any number of jumps, $M \in [0, \infty)$. We can then write the probability of the trajectory x as

$$P_{\beta}(\mathbf{x}) = \frac{1}{Z_{\beta}} \delta[x(0), x(\beta)] e^{-\int_{0}^{\beta} dt \, H_{x(t)}^{c}} \prod_{m=1}^{\mathcal{K}(\mathbf{x})} K_{x_{m-1} \to x_{m}}.$$
 (7)

The probability Eq. (7) can be related to that for trajectories generated by a continuous-time Markov dynamics in the following way. We define the Markov generator $\hat{W} = \hat{K} - \hat{R}$, with the transitions given by \hat{K} and their associated escape rates, \hat{R} ,

$$\hat{R} = \sum_{x,y \neq x} K_{x \to y} |x\rangle \langle x| = \sum_{x} R_{x} |x\rangle \langle x|.$$
(8)

Since we consider quantum systems without a sign problem [12,13], we can assume that the probabilities Eq. (7) are positive real numbers. With these definitions, the relation between the Hamiltonian and the stochastic generator is

$$\hat{H} = -(\hat{W} + \hat{R} - \hat{H}^{c}).$$
(9)

It follows that the trajectories that define the partition function Eq. (6) are related to those generated by the stochastic dynamics of \hat{W} with the following properties:

(1) The trajectory must start and finish in the same configuration, i.e., it should be a *stochastic bridge*. A trajectory which does not meet this criterion has zero probability of occurring. (2) The probability of a trajectory that satisfies (i) is exponentially biased (with respect to the probability of occurring under \hat{W}) by the time integral of $\hat{H}^c - \hat{R}$ along the trajectory. That is, its probability is multiplied by the factor $e^{-\int_0^{\beta} d\tau (H_{x(t)}^c - R_{x(t)})}$.

Condition (i) is needed due to the imposition of periodic paths in Eq. (4). The bias in (ii) accounts for the difference in the diagonal parts of \hat{W} and \hat{H} . Many methods have been developed to efficiently sample biased dynamics as in (ii). One popular approach is TPS, see, e.g., Refs. [19,35,36]. Here, we will use TPS, focusing on techniques for sampling trajectory updates which are local in space, cf. Refs. [17–19], and which also account for (i).

B. Calculating observables

The thermal expectation value of an observable \hat{O} is given by

$$\langle \hat{O} \rangle_{\beta} = \frac{\sum_{x} \langle x | \hat{O} e^{-\beta \hat{H}} | x \rangle}{Z_{\beta}}.$$
 (10)

This can be rewritten as an imaginary-time average by noticing that we can arbitrarily move \hat{O} through the exponential due to the properties of the trace:

$$\langle \hat{O} \rangle_{\beta} = \frac{1}{\beta} \int_{0}^{\beta} dt \, \frac{\sum_{x} \langle x | e^{-(\beta - t)\hat{H}} \hat{O} e^{-t\hat{H}} | x \rangle}{Z_{\beta}}.$$
 (11)

To consider how to compute Eq. (11) from trajectories, we will deal separately with the cases when \hat{O} is diagonal and off-diagonal.

For the case of a diagonal operator, $\hat{O}_{\text{diag}} = \sum_{x} O(x) |x\rangle \langle x|$, Eq. (11) can be directly written as a sum over all trajectories with the probability given by Eq. (7). This leads to

$$\langle \hat{O}_{\text{diag}} \rangle_{\beta} = \beta^{-1} \sum_{\{\boldsymbol{x}\}} P_{\beta}(\boldsymbol{x}) \mathcal{O}(\boldsymbol{x}),$$
 (12)

where $\mathcal{O}(\mathbf{x}) = \int_0^\beta dt \ O(\mathbf{x}(t))$ is the time-integrated trajectory observable. That is, we average the value of the observable over all times and over all trajectories.

For an off-diagonal operator, $\hat{O}_{off-diag} = \sum_{y \neq x} O_{x \rightarrow y} |y\rangle \langle x|$, we can expand both exponentials in Eq. (11). This expansion in the integrand of Eq. (11) gives a sum over all trajectories which have the transition $x \rightarrow y$ at the time *t*, summed over $y \neq x$. Compared to Eq. (6), the jumps at time *t* appear with factors $O_{x \rightarrow y}$ rather than $K_{x \rightarrow y}$. Thus, if we want to express $\langle \hat{O}_{off-diag} \rangle$ as a sum over trajectories with the probability given by Eq. (7), we need to account for the change $K_{x \rightarrow y} \rightarrow O_{x \rightarrow y}$ at time *t*. Furthermore, since we are time averaging, we can replace the integration over time by $\mathcal{K}_{x \rightarrow y}(x)$ which counts the number of jumps $x \rightarrow y$ which occur in trajectory *x*. We can then write

$$\langle \hat{O}_{\text{off-diag}} \rangle_{\beta} = \beta^{-1} \sum_{\{x\}} P_{\beta}(x) \sum_{x \neq y} \frac{\mathcal{K}_{x \to y}(x)}{K_{x \to y}} O_{x \to y}.$$
 (13)

III. OPTIMAL SAMPLING

We now explain how trajectories can be sampled efficiently and exactly from the partition function Eq. (6) by means of a rejection-free form of TPS. A general method for converging to an ensemble of trajectories such as Eq. (7) is TPS, a form of Monte Carlo sampling in trajectory space [35]. A typical problem with standard TPS is that acceptance of trajectory updates can become exponentially small in system size (and/or the length of the trajectory), thus slowing down convergence.

In the language of stochastic dynamics, the procedure we now describe is sometimes referred to as obtaining the *Doob dynamics* [14–16], a proper (normalized) stochastic dynamics derived from the original generator \hat{W} , which generates trajectories with a conditioned/biased probability $P_{\beta}(\mathbf{x})$, such as Eq. (7). While this approach is exact in theory, it often relies on the computation of expressions which are, in practice, analytically (and often numerically) intractable. However, we will show later in the paper that we can in fact implement this general idea via optimal local updates.

A. Edge configurations

The first obstacle is to sample the initial/final configuration of trajectories, x_0 . While Eq. (6) has no explicit probability distribution for the initial configuration x_0 , its probability will be decided by its possible transitions and its diagonal component in Eq. (1). That is, we choose some initial configuration, x_0 , with probability

$$P(x_0) = \frac{\langle x_0 | e^{-\beta H} | x_0 \rangle}{Z_\beta}.$$
 (14)

Once the trajectory edges have been selected, we can then sample the remainder of the trajectory from the subset of trajectories which have the required boundary conditions. That is, we want to sample from the reduced dynamics

$$Z_{\beta}(x_{i}, x_{f}) = \langle x_{f} | e^{-\beta H} | x_{i} \rangle, \qquad (15)$$

where x_i and x_f are the initial and final configurations. Note that, although we are only interested in trajectories which have $x_i = x_f$, it will be useful to solve the more general problem later for the TFIM and QTPM.

B. Continuous-time dynamics

The continuous-time Monte Carlo (CTMC) method [37–39] is a dynamical protocol which can be employed to sample the dynamics of Eq. (6). While this method is straightforward for a normalized time-homogeneous dynamics, it will be useful to formulate the general case of time-dependent dynamics for what follows.

Consider time-dependent instantaneous transition rates, $\tilde{K}_{x \to y}(t; \beta, x_f) \ge 0$, for transitions from configuration *x* to configuration *y* at time $t \le \beta$. We use the tilde to distinguish these from the off-diagonal entries of the Hamiltonian, cf. Eq. (2), and also allow for a dependence on the final configuration, x_f , and the overall trajectory time, β . The associated escape rate from configuration *x* at time *t* is

$$\tilde{R}_{x}(t;\beta,x_{\rm f}) = \sum_{y \neq x} \tilde{K}_{x \to y}(t;\beta,x_{\rm f}).$$
(16)

The escape rate, in turn, determines the distribution for the waiting time τ , P_x , for a jump out of x after t:

$$P_{x}(\tau;t,\beta,x_{\rm f}) = R_{x}(t+\tau;\beta,x_{\rm f})$$
$$\times \exp\left[-\int_{t}^{t+\tau} dt' \tilde{R}_{x}(t';\beta,x_{\rm f})\right]. \quad (17)$$

After a waiting time has been drawn from Eq. (17), the transition into y at time $t + \tau$ is chosen with probability

$$P_x(y \mid t + \tau; \beta, x_f) = \frac{\tilde{K}_{x \to y}(t + \tau; \beta, x_f)}{\sum_{z \neq x} \tilde{K}_{x \to z}(t + \tau; \beta, x_f)}.$$
 (18)

C. Optimal transition rates

Given a dynamics defined by transition rates $\tilde{K}_{x \to y}(t; \beta, x_f)$, the CTMC method generates trajectories with probability density

$$\tilde{P}_{\beta}(\mathbf{x}) = \delta[x_{i}, x(0)] \ e^{-\int_{0}^{r} dt \ R_{x(t)}(t;\beta,x_{f})} \\ \times \prod_{m=1}^{\mathcal{K}(\mathbf{x})} \tilde{K}_{x_{m-1} \to x_{m}}(t_{m};\beta,x_{f})$$
(19)

for trajectories starting from x_i (where x_f acts as a parameter in the definition of the rates). However, we are interested in sampling trajectories from the (conditioned and tilted) distribution Eq. (7) where the final configuration is x_f , that is,

$$P_{\beta}(\mathbf{x}) = \frac{\delta[x(\beta), x_{\rm f}]\delta[x(0), x_{\rm i}]}{Z_{\beta}(x_{\rm i}, x_{\rm f})} \ e^{-\int_{0}^{\beta} dt H_{x(t)}^{c}} \\ \times \prod_{m=1}^{\mathcal{K}(\mathbf{x})} K_{x_{m-1} \to x_{m}}.$$
(20)

In the expression above, we allow (for later convenience) the final configuration to be fixed to a different value from the initial one. As written, Eq. (20) is not the distribution of trajectories generated with transition rates $K_{x_{i-1} \rightarrow x_i}$ due to Eq. (20) satisfying conditions (i) and (ii) above (the former generalized to some fixed initial and final configurations), and the need for the explicit normalization $Z_\beta(x_i, x_f)$. Furthermore, while Eq. (19) appears simpler than Eq. (20) due to the apparent absence of the conditioning of the boundary conditions in time, this is accounted for by the fact that the transition rates in Eq. (19) are *time dependent*.

Our aim is to find a dynamics $\tilde{K}_{x \to y}(t; \beta, x_f)$ such that $\tilde{P}_{\beta}(\mathbf{x}) = P_{\beta}(\mathbf{x})$; that is, the optimal dynamics (or *Doob dynamics*) for sampling $P_{\beta}(\mathbf{x})$ [15]. These optimal transition rates can be computed from the probability that the transition $x \to y$ occurs at the time t under the original dynamics, conditioned by the probability that the trajectory is in configuration x at time t,

$$\tilde{K}_{x \to y}(t; \beta, x_{f}) = \frac{\langle x_{f} | e^{-(\beta - t)\hat{H}} \hat{K}_{x \to y} e^{-t\hat{H}} | x_{i} \rangle}{\langle x_{f} | e^{-(\beta - t)\hat{H}} | x \rangle \langle x | e^{-t\hat{H}} | x_{i} \rangle}$$

$$= \frac{\langle x_{f} | e^{-(\beta - t)\hat{H}} | y \rangle}{\langle x_{f} | e^{-(\beta - t)\hat{H}} | x \rangle} K_{x \to y}$$

$$= \frac{Z_{\beta - t}(y, x_{f})}{Z_{\beta - t}(x, x_{f})} K_{x \to y}, \qquad (21)$$

where in the second line we have inserted the definition of $\hat{K}_{x \to y}$, Eq. (2), and in the third line we used Eq. (15). From this, we are able to calculate the escape rate from the configuration *x* at time *t*,

$$\tilde{R}_{x}(t;\beta,x_{\rm f}) = \sum_{y} \tilde{K}_{x \to y}(t;\beta,x_{\rm f})$$

$$= \frac{\langle x_{\rm f} | e^{-(\beta-t)\hat{H}} \hat{K} | x \rangle}{\langle x_{\rm f} | e^{-(\beta-t)\hat{H}} | x \rangle}$$

$$= H_{x}^{\rm c} - \frac{\partial}{\partial t} \ln Z_{\beta-t}(x,x_{\rm f}), \qquad (22)$$

where we have used Eq. (2) to account for all possible transitions out of x and the derivative of Eq. (15) to obtain the last line. Based on the rates from Eqs. (21) and (22), this time-dependent dynamics gives $\tilde{P}_{\beta}(\mathbf{x}) = P_{\beta}(\mathbf{x})$ if the distribution of initial configurations is set to be x_i [40]. Appendix A explains how the optimal transition rates can be used with the CTMC algorithm from the previous section to sample the time-dependent dynamics.

D. Example: Single two-level system

As an illustration of the ideas above, we consider a simple problem which will be important later in this paper: a single two-level system with the states $\sigma = \{-1, +1\}$, which we will refer to as a spin.

A generic Hamiltonian for such a system reads

$$\hat{H}^{\rm spin} = -\hat{X} - g\hat{Z},\tag{23}$$

where \hat{X} , \hat{Y} , and \hat{Z} are the usual Pauli operators. By calculating the matrix exponential of Eq. (23), is can be shown that the reduced dynamics, Eq. (15), takes the form

$$Z_t(\sigma_i, \sigma_f) = \begin{cases} \cosh(\theta_t) + \frac{g}{\sqrt{1+g^2}} \sinh(\theta_t), & \sigma_{i,f} = 1\\ \cosh(\theta_t) - \frac{g}{\sqrt{1+g^2}} \sinh(\theta_t), & \sigma_{i,f} = -1\\ \frac{1}{\sqrt{1+g^2}} \sinh(\theta_t), & \sigma_i \neq \sigma_f, \end{cases}$$
(24)

where $\theta_t = t\sqrt{1+g^2}$. Using Eq. (24), we are able to define a time-dependent stochastic dynamics which allows us to exactly sample trajectories with probabilities given by Eq. (7). The initial/final configuration σ_0 is chosen with probability $Z_\beta(\sigma_0, \sigma_0)/\sum_{\sigma} Z_\beta(\sigma, \sigma)$, and the time-dependent transition rates for the dynamics are

$$\tilde{K}_{\sigma \to -\sigma}(t; \beta, \sigma_0) = \frac{Z_{\beta - t}(-\sigma, \sigma_0)}{Z_{\beta - t}(\sigma, \sigma_0)}.$$
(25)

By simply running this dynamics, we sample the trajectories we need for computing thermal averages in this problem. The results are shown in Figs. 1(a) and 1(b) for the average magnetizations $\langle \hat{X} \rangle$ and $\langle \hat{Z} \rangle$ for various inverse temperatures β . The data points show the trajectory-averaged values using Eqs. (12) and (13). The numerical results coincide with the exact averages:

$$\langle \hat{Z} \rangle = g \langle \hat{X} \rangle = \frac{g}{\sqrt{1+g^2}} \tanh(\beta \sqrt{1+g^2}).$$
 (26)



FIG. 1. Sampling of a single spin. The magnetizations (a) $\langle \hat{X} \rangle$ and (b) $\langle \hat{Z} \rangle$ for the Hamiltonian Eq. (23), measured from trajectory sampling (data points) and compared to the exact result (solid lines) for various inverse temperatures β .

IV. TRANSVERSE FIELD ISING MODEL IN DIMENSIONS ONE AND TWO

We now show how the trajectory sampling approach can be used to sample from the Boltzmann distribution of the TFIM with N spins,

$$\hat{H} = -h \sum_{i=1}^{N} \hat{X}_i - J \sum_{\langle i,j \rangle} \hat{Z}_i \hat{Z}_j, \qquad (27)$$

where \hat{X}_i and \hat{Z}_i are the Pauli operators acting on site *i*, and nearest neighbors are denoted by $\langle i, j \rangle$. The method we use in this section can be equally adapted to accommodate for any other potential (the second summation) with the same single-body kinetic term, see Appendix B for an example with a global diagonal operator. In the next section, we demonstrate how to deal with other kinetic terms.

In what follows, we set h = 1 and consider periodic boundary conditions in space. Sampling directly from the full partition function Eq. (4) is difficult. We instead use a singlespin update scheme, whereby we update the entire trajectory for a single spin, keeping the trajectories of all other spins fixed by sampling directly from a reduced partition function.

A. Redrawing trajectories for an individual spin

We now explain how our formulation can be used to perform single spin updates of the many-body trajectories. These updates adjust the entire trajectory of a single spin within the many-body trajectory, modeling the other spins as an effective time-dependent *environment* (or *heat bath* [17]). We stress that this update is equivalent to that presented in Ref. [17], which considered Eq. (27) on the Bethe lattice, and similarly that of Ref. [18] for the classical Ising model. Here, we instead use this approach on the TFIM in 1D and 2D to demonstrate how the update can be used to investigate the ground-state properties and phase transitions of quantum lattice models.

1. Factorization of the partition function

We consider the many-body trajectory $\boldsymbol{\omega}$ as a collection of the *N* individual spin trajectories, $\boldsymbol{\omega} = \{\boldsymbol{\sigma}^1, \cdots, \boldsymbol{\sigma}^N\}$, with $\boldsymbol{\sigma}^j$

denoting the time series of transitions for spin *j*,

$$\boldsymbol{\sigma}^{j} = \left\{ \left(t_{0}^{j}, \sigma_{0}^{j}\right), \left(t_{1}^{j}, \sigma_{1}^{j}\right), \cdots, \left(t_{M_{j}}^{j}, \sigma_{M_{j}}^{j}\right) \right\},$$
(28)

each with a total of M_j transitions. The partition function becomes

$$Z_{\beta}^{\text{TFIM}} = \sum_{\{\omega\}} \left[\prod_{i=1}^{N} \delta[\sigma^{i}(0), \sigma^{i}(\beta)] \right] e^{\int_{0}^{\beta} dt \sum_{\langle i,j \rangle} J\sigma^{i}(t)\sigma^{j}(t)} \\ \times \prod_{j=1}^{N} \prod_{m=1}^{\mathcal{K}(\sigma^{j})} K_{\sigma_{m-1}^{j} \to \sigma_{m}^{j}}.$$
(29)

In the above, all $K_{\sigma \to \sigma'} = 1$, cf. Eq. (27), but we keep them in the expression to keep track of the spin transitions.

At this point, we notice that we are able to factorize Eq. (29) into a factor which depends on a specific σ^l , and a factor which does not,

$$Z_{\beta}^{\text{TFIM}} = \sum_{\{\omega\}} \left(\delta[\sigma^{l}(0), \sigma^{l}(\beta)] e^{\int_{0}^{\beta} dt \sum_{i;(i,l)} J\sigma^{i}(t)\sigma^{l}(t)} \\ \times \prod_{m=1}^{\mathcal{K}(\sigma^{l})} K_{\sigma_{m-1}^{l} \to \sigma_{m}^{l}} \right) \prod_{p \neq l} \left(\delta[\sigma^{p}(0), \sigma^{p}(\beta)] \\ \times e^{\int_{0}^{\beta} dt \frac{1}{2} \sum_{j \neq l;(j,p)} J\sigma^{p}(t)\sigma^{j}(t)} \prod_{n=1}^{\mathcal{K}(\sigma^{p})} K_{\sigma_{n-1}^{p} \to \sigma_{n}^{p}} \right), \quad (30)$$

where $\{i; \langle i, l \rangle\}$ denotes sites *i* which are nearest neighbors of *l*, and $\{j \neq l; \langle j, p \rangle\}$ sites *j* different from *l* which are nearest neighbors of *p*.

We can now express Eq. (30) as a sum over each spin l and for each l over the *partial trajectories*

$$\boldsymbol{\sigma}^{(l)} = \{\boldsymbol{\sigma}^1, \cdots, \boldsymbol{\sigma}^{l-1}, \boldsymbol{\sigma}^{l+1}, \cdots \boldsymbol{\sigma}^N\}, \quad (31)$$

that is, the individual trajectories of all the spins other than l. The first factor in Eq. (30) depends on the trajectory of spin l and its neighboring spins through the interactions $\sum_{i:\langle i,l \rangle} J\sigma^i(t)\sigma^l(t)$, which we rewrite in the form of an effective time-dependent longitudinal field:

$$g^{l}(t) := g(t \mid \boldsymbol{\sigma}^{(l)}) = \sum_{i; \langle i, l \rangle} J \sigma^{i}(t).$$
(32)

The second factor in Eq. (30) only depends on the trajectories of the spins $p \neq l$, and we denote it $\mathcal{N}(\boldsymbol{\sigma}^{(l)})$. The partition function can now be written as

$$Z_{\beta}^{\text{TFIM}} = N^{-1} \sum_{l} \sum_{\{\sigma^{(l)}\}} \mathcal{N}(\sigma^{(l)}) \sum_{\{\sigma^{l}\}} \delta[\sigma^{l}(0), \sigma^{l}(\beta)]$$

$$\times e^{\int_{0}^{\beta} dt \, g^{l}(t) \, \sigma^{l}(t)} \prod_{m=1}^{\mathcal{K}(\sigma^{l})} \mathcal{K}_{\sigma_{m-1}^{l} \to \sigma_{m}^{l}}$$

$$= N^{-1} \sum_{l} \sum_{\{\sigma^{(l)}\}} \mathcal{N}(\sigma^{(l)}) Z_{\beta}^{l}[g^{l}]. \tag{33}$$

Now suppose we have some many-body trajectory $\boldsymbol{\omega}$, and we have randomly chosen to update site *l*. The partition function for spin *l* is $Z_{\beta}^{l}[g^{l}]$, with the time-dependent longitudinal field $g^{l}(t)$ which is piecewise constant, meaning it remains



FIG. 2. Single-spin update scheme for the 1D TFIM. The spin *l* is updated while all the other spins are kept fixed. Only the spins l - 1 and l + 1 (red) contribute to the dynamics of the *l*th spin. These spins change at the times τ_m (dashed grey lines). The update proceeds through the following steps: (a) The trajectory for spin *l* is discarded. (b) We then sample the configurations σ_m (blue circles) at the times in which the effective longitudinal field instantaneously changes, as explained. (c) We then sample bridges between each of the σ_m under a static longitudinal field to construct the full trajectory (blue line).

constant except for a discrete number of times, \tilde{M} , when it instantaneously transitions to some other value. \tilde{M} is the total number of times any of the neighboring spins transition. The partition function can now be written in the following way:

$$Z^{l}_{\beta}[g^{l}] = \sum_{\sigma_{0}, \cdots, \sigma_{\tilde{M}}} \prod_{m=0}^{\tilde{M}} Z_{\Delta \tau_{m}}(\sigma_{m}, \sigma_{m+1}; g^{l}(\tau_{m})).$$
(34)

The spin configurations $\sigma_m = \sigma^l(\tau_m)$ are the configurations of spin *l* at the times τ_m where the field $g^l(t)$ instantaneously changes, and we set $\sigma_{\tilde{M}+1} = \sigma_0$. Equation (34) is a sum over *all* possible spin configurations at each of these times. The weight for each sequence of spin configurations { σ_m } is given by the product in Eq. (34). The factors $Z_{\Delta\tau_m}(\sigma_m, \sigma_{m+1}; g^l(\tau_m))$ are exactly the partition function of the single spin problem, Eq. (24), with time extent $\Delta\tau_m = \tau_{m+1} - \tau_m$ and longitudinal field $g^l(\tau_m)$, which is taken to be the value *after* the *m*th transition of neighboring spins, and can be written as

$$Z_{\Delta\tau_m}(\sigma_m, \sigma_{m+1}; g^l(\tau_m)) = \sum_{\{\sigma_m^l\}} \delta[\sigma^l(\tau_m), \sigma_m] \delta[\sigma^l(\tau_{m+1}), \sigma_{m+1}] \times e^{g^l(\tau_m) \int_{\tau_m}^{\tau_{m+1}} dt \, \sigma^l(t)} \prod_{k_m=1}^{\mathcal{K}(\sigma_m^l)} K_{\sigma_{k_m-1}^l \to \sigma_{k_m}^l}.$$
 (35)

In the above equation, $\sigma_m^l = \sigma^l(\tau_m : \tau_{m+1})$ denotes the partial trajectory of spin σ^l between times τ_m and τ_{m+1} . It is simple to check that Eq. (34) with Eq. (35) gives $Z_{\beta}^l[g^l]$.

We now explain the strategy to update a trajectory ω , making use of the above equations and Fig. 2, which demonstrates the update for the 1D TFIM. We first randomly select a spin $l \in \{1, ..., N\}$ to update, keeping the trajectories for all other spins fixed. The trajectories of spins neighboring *l* define the time-dependent longitudinal field $g^l(t)$. The process of selecting spin *l* and determining the time-dependent field takes us from Eq. (29) to Eq. (33), and allows us to sample a trajectory for spin *l* through the partition function $Z^l_{\beta}[g^l]$. We then exploit the fact that the longitudinal field is constant in between the transition times τ_m for spins neighboring *l*. In particular, we write $Z^l_{\beta}[g^l]$ as a sum over all *edge configurations*, σ_m , which gives Eq. (34). The edge configurations can be sampled using transfer matrices, which is explained in Sec. IV A 2, see Fig. 2(b). Finally, we sample *stochastic bridges* between configurations σ_m and σ_{m+1} as explained in Sec. IV A 3, see Fig. 2(c).

2. Sampling the edge configurations

Given some time-dependent longitudinal field $g^l(t)$, we wish to sample the states of the spin $\sigma_m = \sigma^l(\tau_m)$ at the times τ_m which the longitudinal field changes value. The probability of observing the sequence of configurations $\{\sigma_0, \dots, \sigma_{\tilde{M}+1}\}$ is

$$P(\sigma_0, \cdots, \sigma_{\tilde{M}+1})$$

$$= \frac{1}{Z_{\beta}^l[g^l]} \prod_{m=0}^{\tilde{M}} Z_{\Delta \tau_m}(\sigma_m, \sigma_{m+1}; g^l(\tau_m)).$$
(36)

Each of the partition functions $Z_{\Delta\tau_m}(\sigma_m, \sigma_{m+1}; g^l(\tau_m))$ has four possible values. Note that if we were to think of each σ_m as a spin on a periodic 1D lattice with $\tilde{M} + 1$ lattice sites, then $Z_{\Delta\tau_m}(\sigma_m, \sigma_{m+1}; g^l(\tau_m))$ can be thought of as a *transfer matrix* between lattice sites *m* and *m* + 1.

To efficiently sample the sequence of configurations, we start by sampling σ_0 . The probability of observing some σ_0 is given by

$$P(\sigma_0) = \frac{1}{Z_{\beta}^l[g^l]} \sum_{\sigma_1, \dots, \sigma_{\bar{M}}} \prod_{m=0}^{\tilde{M}} Z_{\Delta \tau_m}(\sigma_m, \sigma_{m+1}; g^l(\tau_m)), \quad (37)$$

with $\sigma_{\tilde{M}+1} = \sigma_0$. Note that calculating $P(\sigma_0)$ is equivalent to performing $\tilde{M} + 1$ matrix multiplications (for 2 × 2 matrices), and can then be calculated numerically. The configurations σ_m for $m = 1, \ldots, \tilde{M}$ can be calculated recursively using

$$P(\sigma_{m} | \{\sigma_{0}, \cdots, \sigma_{m-1}\}) = \frac{\sum_{\sigma_{m+1}, \cdots, \sigma_{\tilde{M}}} \prod_{m=0}^{\tilde{M}} Z_{\Delta \tau_{m}}(\sigma_{m}, \sigma_{m+1}; g^{l}(\tau_{m}))}{P(\sigma_{m-1} | \{\sigma_{0}, \cdots, \sigma_{m-2}\})}.$$
 (38)

The numerator is calculated as a multiplication over $\tilde{M} + 1 - m$ matrices and *m* real numbers, and the denominator is known from the previous iteration of the calculation. It is simple to show that sampling from each of these $\tilde{M} + 1$ distributions

yields a sequence of configurations sampled with probability given by Eq. (36).

3. Sampling the full trajectory

After we have sampled the configurations σ_m , all that is left to do is to sample from each of the partition functions $Z_{\Delta \tau_m}(\sigma_m, \sigma_{m+1}; g^l(\tau_m))$. That is, we must sample stochastic bridges between configurations σ_m and σ_{m+1} as was done previously for the two-level system. This indicates that the trajectory between times τ_m and τ_{m+1} is sampled from a dynamics initialized in configuration σ_m , with transition rates

$$\tilde{K}_{\sigma^{l} \to -\sigma^{l}}(t; \Delta \tau_{m}, \sigma_{m+1}) = \frac{Z_{\Delta \tau_{m}-t}(-\sigma^{l}, \sigma_{m+1}; g^{l}(\tau_{m}))}{Z_{\Delta \tau_{m}-t}(\sigma^{l}, \sigma_{m+1}; g^{l}(\tau_{m}))}.$$
 (39)

Each partial trajectory between times τ_m and τ_{m+1} is sampled with probability

$$P(\boldsymbol{\sigma}_{m}^{l} \mid \boldsymbol{\sigma}^{(l)}) = \frac{\delta[\sigma_{m}, \sigma^{l}(\tau_{m})] \,\delta[\sigma_{m+1}, \sigma^{l}(\tau_{m+1})]}{Z_{\Delta \tau_{m}}(\sigma_{m}, \sigma_{m+1}; g^{l}(\tau_{m}))} \times e^{g(\tau_{m}) \int_{\tau_{m}}^{\tau_{m+1}} dt \,\sigma^{l}(t)} \prod_{k_{m}=1}^{\mathcal{K}(\sigma_{k}^{l})} K_{\sigma_{k_{m}-1}^{l} \to \sigma_{k_{m}}^{l}}.$$
 (40)

Concatenating the partial trajectories gives the fully sampled trajectory for the spin σ^{l} , which is sampled with probability

$$P(\boldsymbol{\sigma}^{l}|\boldsymbol{\sigma}^{(l)}) = \frac{1}{Z_{\beta}[g^{l}]} \prod_{m=0}^{\tilde{M}} \left[Z_{\Delta\tau_{m}}(\sigma_{m}, \sigma_{m+1}; g^{l}(\tau_{m})) P(\boldsymbol{\sigma}_{m}^{l}|\boldsymbol{\sigma}^{(l)}) \right].$$

$$(41)$$

Note that the trajectories are defined to be piecewise continuous, and so we can drop the intermediate $\delta[\sigma_m, \sigma(\tau_m)]$ terms in Eq. (40):

$$P(\boldsymbol{\sigma}^{l} \mid \boldsymbol{\sigma}^{(l)}) = \frac{\delta[\sigma(\beta), \sigma(0)]}{Z_{\beta}[g^{l}]} e^{\int_{0}^{\beta} dt \sum_{i:\langle i,l \rangle} J\sigma^{i}(t)\sigma(t)} \times \prod_{k=1}^{\mathcal{K}(\boldsymbol{\sigma})} K_{\sigma_{k-1} \to \sigma_{k}}.$$
(42)

4. Detailed balance

For the trajectory update to be rejection-free, it must obey detailed balance with respect to Eq. (29). It can be verified that this is true by considering the transition probabilities. That is, given some trajectory $\boldsymbol{\omega}$, the probability to generate a new trajectory $\tilde{\boldsymbol{\omega}}$ is

$$P(\tilde{\boldsymbol{\omega}} \mid \boldsymbol{\omega}) = N^{-1} P(\tilde{\boldsymbol{\sigma}}^l \mid \boldsymbol{\sigma}^{(l)}), \tag{43}$$

where the factor of *N* comes from the fact that site *l* is chosen randomly. Note that since $P(\sigma^{l} | \sigma^{(l)}) \propto P_{\beta}(\omega)$, it follows that

$$\frac{P(\tilde{\boldsymbol{\omega}} \mid \boldsymbol{\omega})}{P(\boldsymbol{\omega} \mid \tilde{\boldsymbol{\omega}})} = \frac{P(\tilde{\boldsymbol{\sigma}}^{l} \mid \boldsymbol{\sigma}^{(l)})}{P(\boldsymbol{\sigma}^{l} \mid \boldsymbol{\sigma}^{(l)})} = \frac{P_{\beta}(\tilde{\boldsymbol{\omega}})}{P_{\beta}(\boldsymbol{\omega})}$$
(44)

and detailed balance is obeyed.

B. Monte Carlo method

The method of resampling a trajectory for a single spin given in the previous section can now be used to implement a Monte Carlo algorithm to sample the many-body trajectories using the following steps:

(1) Create some initial seed trajectory, $\boldsymbol{\omega} = (\boldsymbol{\sigma}^1, \cdots, \boldsymbol{\sigma}^N)$ with inverse temperature β .

- (2) Choose a random lattice site $l \in [1, N]$.
- (3) Generate a new trajectory, σ^l , for the spin.
- (4) Repeat from step 2 until convergence.

There are various ways to generate an initial seed trajectory. The key requirement is periodicity in the time interval $[0, \beta]$. The simplest way to achieve this is to generate the trajectory for each spin independently with a noninteracting dynamics, but allowing for the possibility of a local longitudinal field (see Sec. III D). The choice of the initial trajectory can be guided by the phase one is trying to target. For example, for the TFIM with J/h > 1, we expect to see ferromagnetic behavior, and so we could choose a large local longitudinal field to force an initial trajectory with a higher likelihood of magnetic ordering. A second approach, which we discuss in Sec. V B, is *thermal annealing*.

For |J/h| > 1, there is spontaneous breaking of the symmetry at the level of the ground state. Indeed, depending on the initial trajectory seed, the single-spin update approach will break the \mathbb{Z}_2 symmetry and will only be ergodic over one of the ground states. In practice, one should randomly perform the global update $\sigma^i(t) \rightarrow -\sigma^i(t)$ for all spins simultaneously to ensure both states are explored. To investigate the effects of the local updates, we choose not to do this here.

C. Continuous phase transition in the TFIM

We now demonstrate the effectiveness of our approach for the 1D and 2D TFIM. For fixed h = 1, both models are known to undergo a continuous phase transition at the critical points $J_c = 1$ [41] and $J_c \approx 0.32747$ [20,21,23], respectively. We investigate their ground state by using the previously described trajectory sampling algorithm with $\beta = 128$ and a variety of system sizes, N. Figures 3(a) and 3(b) show the average transverse magnetization,

$$M_x = N^{-1} \sum_{i=1}^{N} \langle \hat{X}_i \rangle, \tag{45}$$

in 1D and 2D, respectively, while Figs. 3(c) and 3(d) show the two-spin correlator:

$$M_{zz} = N^{-1} \sum_{\langle i,j \rangle} \langle \hat{Z}_i \hat{Z}_j \rangle.$$
(46)

For 1D, we compare our results for systems of size N = 16 to 1024 to those from infinite matrix product states (iMPS) [42,43]. These methods allow us to determine the properties of the 1D chain to high accuracy. In particular, we also use the infinite time-evolving block decimation method [42,43]. For 2D, we show results for systems $N = 4 \times 4$ to 32×32 . These demonstrate agreement with the 2D density matrix renormalization group (DMRG) [44–46] for $N = 4 \times 4$ (dashed line).

The characterization of a continuous phase transition is facilitated through the use of an order parameter which displays a discontinuous derivative at the critical point J_c in the large system-size limit. The commonly used order parameter for



FIG. 3. Ground-state properties of the TFIM. Numerical results using the single-spin update for the 1D TFIM (top) for system sizes N = 16 to 1024 and the 2D TFIM (bottom) with $N = L^2$ and L = 4 to 32. We show the average transverse magnetization, M_x , in (a) and (b); the average two-spin correlator, M_{zz} , in (c) and (d); and the average magnetization, $M_{|z|}$, in (e) and (f). Data points are the results from trajectory sampling. The dashed black line in the top row of panels is from the iMPS, and from the 2D-DMRG with L = 4 for the bottom row. The value of J_c in 1D is $J_c = 1$ and $J_c \approx 0.32747$ in 2D.

Ising models is the longitudinal magnetization,

$$M_{|z|} = N^{-1} \left\langle \left| \sum_{i=1}^{N} \hat{Z}_{i} \right| \right\rangle.$$
(47)

Figures 3(e) and 3(f) demonstrate a sharp increase for $M_{|z|}$ at the critical point, $J/J_c = 1$.

D. Trajectory autocorrelations

We now investigate the autocorrelation properties of the sampling dynamics using the single-spin update for the 1D and 2D TFIM. Under TPS, we generate a *Markov chain of trajectories*, $\{\omega_0 \rightarrow \omega_1 \rightarrow \cdots \rightarrow \omega_{N_{traj}}\}$. To test the convergence to the target trajectory ensemble, cf. Eq. (7), we define the autocorrelation between two trajectories in the Markov chain separated by μ TPS updates,

$$\mathcal{C}(\boldsymbol{\omega}_{\nu},\boldsymbol{\omega}_{\nu+\mu}) = \frac{1}{\beta N} \sum_{j=1}^{N} \int_{0}^{\beta} d\tau \, \sigma_{\nu}^{j}(\tau) \, \sigma_{\nu+\mu}^{j}(\tau), \qquad (48)$$

where $\sigma_{\nu}^{j}(\tau)$ indicates the value of the *j* spin of the ν th trajectory at time τ . The trajectory ensemble average can then be estimated from the Markov chain,

$$C_{\mu} = \frac{1}{N_{\text{traj}} - \mu} \sum_{\nu=1}^{N_{\text{traj}} - \mu} \frac{\mathcal{C}(\boldsymbol{\omega}_{\nu}, \boldsymbol{\omega}_{\nu+\mu}) - \langle \hat{Z} \rangle^2}{1 - \langle \hat{Z} \rangle^2}, \qquad (49)$$

where $\langle \hat{Z} \rangle$ is the equilibrium expectation value of \hat{Z}_i (for any *i*). With the above definition, this autocorrelator is normalized to be $C_0 = 1$ and $C_{\infty} = 0$ for $N_{\text{traj}} \to \infty$.

We show C_{μ} as a function of J/J_c in Fig. 4 for $\mu/N \in [1, 7]$, for (a) the 1D TFIM and (b) the 2D TFIM. The same data is shown in Figs. 4(c) and 4(d) as a function of μ/N . Far



FIG. 4. Convergence in the TFIM. The autocorrelation between trajectories C_{μ} as a function of J/J_c and $\mu/N \in [1, 7]$ for (a) the 1D TFIM with $N = \beta = 128$ and (b) the 2D TFIM with $N = L^2$, L = 32 and $\beta = 128$. (c), (d) The same data but as a function of μ/N for various values of J/J_c . The black dashed lines show exponential decay.

from the critical point, the trajectory autocorrelation function decays approximately exponentially. Near the critical point, this decay indicates that trajectories remain correlated even after a considerable number of TPS iterations. This is a manifestation of the expected slowdown of Monte Carlo sampling near criticality at the level of trajectory sampling. This is a shortcoming of the method when compared to methods which use nonlocal updates, such as rejection-free cluster updates [20,21] and Worm algorithms [7,8,23]. However, it is important to note that it is not always easy to find cluster updates for interaction terms which are more complex than the Ising interaction in Eq. (27). On the contrary, the approach described here can be generalized to investigate arbitrary diagonal terms with a transverse field, see Ref. [24] for triangular plaquette interactions and Appendix B for an example with a global diagonal operator.

V. QUANTUM TRIANGULAR PLAQUETTE MODEL

We now show how the local update scheme can be generalized to models with more complex kinetic terms in their Hamiltonians. As a concrete example, we consider the QTPM [19,24–29], a generalization of the classical TPM studied in the context of glassy systems [30-32]. This is a model defined on a triangular lattice with interactions between a subset of the triangular plaquettes and a magnetic field transverse to them. We write the Hamiltonian of the QTPM as

$$\hat{H} = -h \sum_{\{i,j,k\} \in \Delta} \hat{X}_i \hat{X}_j \hat{X}_k - J \sum_{i=1}^N \hat{Z}_i,$$
(50)

where \triangle indicates the upward pointing triangular plaquettes in the triangular lattice; see Fig. 5. We have chosen a representation of the QTPM where the interactions are off-diagonal in the classical basis and the magnetic field is longitudinal. This corresponds to the *dual* model of the usual QTPM.

Most numerical studies [19,24,25] indicate that the QTPM has a first-order quantum phase transition at $|J_c/h_c| = 1$ between two distinct phases. In our recent paper, Ref. [24], we used the method presented in Sec. IV with single-spin updates on the Hamiltonian dual to Eq. (50). The aim of this section is to demonstrate how the update scheme can be adjusted to account for Hamiltonians like Eq. (50), using local updates which redraw the trajectories for multiple spins, rather than a single spin.

A. Plaquette updates

The obvious Monte Carlo update is to randomly select a *plaquette* labeled by \triangle encompassing three sites $\{i, j, k\} \in \triangle$. The trajectory of this plaquette is $\eta^{\triangle} = \{\sigma^i, \sigma^j, \sigma^k\}$. The corresponding transitions are those where the three spins in the plaquette flip simultaneously, in accordance with the kinetic term in Eq. (50). Given the eight different possible states of these three spins, it might appear that one needs to consider this eight-level system in the simulations. However, as we will show below, what actually matters is the change in sign in the plaquette magnetization,

$$M_{\wedge} = \sigma^i + \sigma^j + \sigma^k, \tag{51}$$



FIG. 5. Elementary transitions in the TPM. A visualisation of the allowed transitions in the TPM. A plaquette of three spins (spins are shown by blue and red arrows), \triangle , can flip. On the level of the individual plaquette, the plaquette can be in four different disconnected sectors, each containing two configurations and behaving like a two-level system. Within each sector, the system can move between the magnetizations $M_{\triangle} \leftrightarrow -M_{\triangle}$. The plaquettes in the green (left) box show plaquettes with total magnetization $M_{\Delta} = \pm 3$. The three orange boxes (right) show plaquettes with total magnetization M_{\wedge} = ± 1 . The flipping of a neighboring plaquette, Δ' , causes a single spin within \triangle to flip, moving \triangle into a different sector, as demonstrated by the arrows connecting boxes. If it moves between the $M_{\triangle} = |3|$ and $M_{\Delta} = |1|$ sectors, then the sign of the magnetization M_{Δ} is preserved; otherwise, the sign is flipped.

which reduces again the analysis to that of a local two-level system, see Fig. 5.

The time-dependent effective dynamics for the plaquette whose trajectory we choose to update is given by the six neighboring plaquettes which contain any of the spins i, j, or k. Generalizing the approach used for the TFIM above, we do not modify the trajectories of these neighboring plaquettes when updating plaquette \triangle . The neighboring plaquettes will have a total of \tilde{M} flips at times $0 \leq \tau_1 \leq \tau_2 \leq \cdots \leq \tau_{\tilde{M}} \leq \beta$. These transitions will force a single spin in η^{Δ} to flip at each of these times.

1. Factorization of the partition function

We start by writing the partition function as a sum over all possible imaginary-time trajectories [47],

$$Z_{\beta}^{\text{TPM}} = \sum_{\{\omega\}} e^{\sum_{i} \int_{0}^{\beta} dt \, J\sigma^{i}(t)} \prod_{i=1}^{N} \delta[\sigma^{i}(0), \sigma^{i}(\beta)]$$
$$\times \prod_{\Delta} \prod_{m=1}^{\mathcal{K}(\eta^{\Delta})} K_{\eta_{m-1}^{\Delta} \to \eta_{m}^{\Delta}}, \tag{52}$$

where a plaquette trajectory is given by

$$\boldsymbol{y}^{\boldsymbol{\triangle}} = \{(t_0^{\boldsymbol{\triangle}}, \eta_0^{\boldsymbol{\triangle}}), (t_1^{\boldsymbol{\triangle}}, \eta_1^{\boldsymbol{\triangle}}), \cdots, (t_M^{\boldsymbol{\triangle}}, \eta_M^{\boldsymbol{\triangle}})\},$$
(53)

and $K_{\eta_{m-1}^{\triangle} \to \eta_m^{\triangle}}$ denotes the corresponding plaquette flip. As in the case of the TFIM, we can write Eq. (52) in a factorized form. Here, however, there will be three contributions. The first will contain all degrees of freedom and transitions within the plaquette \triangle . The second will contain all the transitions associated with the six other plaquettes that contain any of the spins $\{i, j, k\} \in \Delta$, and it will determine the effective dynamics for the plaquette Δ . The

third factor contains all other contributions which do not affect the dynamics of plaquette \triangle . The partition function reads

$$Z_{\beta}^{\text{TPM}} = \sum_{\{\omega\}} \left[\prod_{i \in \Delta} (\delta[\sigma^{i}(0), \sigma^{i}(\beta)] e^{\int_{0}^{\beta} dt J \sigma^{i}(t)}) \prod_{m=1}^{\mathcal{K}(\eta^{\Delta})} K_{\eta_{m-1}^{\Delta} \to \eta_{m}^{\Delta}} \right] \left[\prod_{\Delta'} \prod_{m'=1}^{\mathcal{K}(\eta^{\Delta'})} K_{\eta_{m'-1}^{\Delta'} \to \eta_{m'}^{\Delta'}} \right] \\ \times \left[\prod_{i \notin \Delta} (\delta[\sigma^{i}(0), \sigma^{i}(\beta)] e^{\int_{0}^{\beta} dt J \sigma^{i}(t)}) \prod_{\Delta''} \prod_{m'=1}^{\mathcal{K}(\eta^{\Delta''})} K_{\eta_{m'-1}^{\Delta''} \to \eta_{m'}^{\Delta''}} \right],$$
(54)

where \triangle' labels the plaquettes which share a site with \triangle , and \triangle'' labels those which do not. We will now single out and propose a trajectory update for just one of the Nplaquettes. The last factor does not depend on any of the spins within plaquette \triangle : as before, we denote $\sigma^{(\triangle)}$ the trajectory of all the spins except from those $\{i, j, k\} \in \Delta$ and write this factor as $\mathcal{N}_{\Delta}(\boldsymbol{\sigma}^{(\Delta)})$. The second factor does depend on the spins within plaquette \triangle , but corresponds to plaquette flips which are not for plaquette \triangle . By this, we mean each lattice site $\{i, j, k\} \in \Delta$ will belong to two other plaquettes. When either of these plaquettes flip, so will the spin $\sigma^{\{i,j,k\}}$. However, the trajectory of each of these plaquettes is fixed, and thus we call this factor $\mathcal{M}_{\Delta}(\{\eta^{\Delta'}\})$, where $\{\eta^{\Delta'}\}$ denotes the trajectories of all neighboring plaquettes. The first factor corresponds to plaquette $\boldsymbol{\bigtriangleup}$ and depends on $\{\eta^{\Delta'}\}$. We can now rewrite Z_{β}^{TPM} as a sum over all plaquettes \triangle and over all trajectories of plaquettes $\{\eta^{(\triangle)}\}$ which are not \triangle :

$$Z_{\beta}^{\text{TPM}} = N^{-1} \sum_{\Delta} \sum_{\{\boldsymbol{\eta}^{(\Delta)}\}} \mathcal{N}_{\Delta}(\boldsymbol{\sigma}^{(\Delta)}) \mathcal{M}_{\Delta}(\{\boldsymbol{\eta}^{\Delta'}\}) Z_{\beta}^{\Delta}[\{\boldsymbol{\eta}^{\Delta'}\}].$$
(55)

The partition function $Z_{\beta}^{\Delta}[\{\eta^{\Delta'}\}]$ is the sum over all possible trajectories for the plaquette Δ , subject to the transitions which occur in the neighboring plaquettes at fixed times.

As was done for the TFIM, we can now factorize the partition function into $\tilde{M} + 1$ components, where \tilde{M} is the total number of times any of the neighboring plaquettes flip. We can write

$$Z_{\beta}^{\Delta}[\{\boldsymbol{\eta}^{\Delta'}\}] = \sum_{\eta_0,\cdots,\eta_{\bar{M}}} \prod_{m=0}^{\bar{M}} Z_{\Delta\tau_m}^{\Delta}(\eta'_m,\eta_{m+1}), \qquad (56)$$

where $\Delta \tau_m = \tau_{m+1} - \tau_m$. The plaquette configuration η'_m is related to η_m by a single spin flip, which is predetermined by the neighboring plaquettes. This is true for all $m \neq 0$: periodic boundary conditions in imaginary time gives the condition $\eta'_0 = \eta_0 = \eta_{\tilde{M}+1}$.

2. Determining the trajectory sector and sampling the edge configurations

An important point to notice is that the trajectory space of η^{Δ} , for some chosen Δ , is composed of four disconnected sectors, which we refer to as the *trajectory sectors*. An initial

plaquette configuration $\eta_0^{\Delta} = \{\sigma_0^i, \sigma_0^j, \sigma_0^k\}$ will belong to one of the four local sectors shown in Fig. 5. At the times τ_m when there is a transition in a neighboring plaquette Δ' , the sector of the plaquette Δ will transition to one of the other three sectors, as shown in Fig. 5; however, the sector it moves to is entirely determined by which plaquette Δ' transitioned. Thus, the initial sector for η_0 will predetermine which local sector the plaquette will occupy at any time.

This observation implies that the local plaquette effectively behaves as a two-level (rather than an eight-level) system, and the flipping of neighboring plaquettes at times t_m changes the energies of the two levels according to the rules shown in Fig. 5. The configurations η_m in Eq. (56) can be sampled using the transfer matrix approach as was done for the TFIM, see Sec. IV A. While the transfer matrices for the QTPM are 8×8 matrices, the previous observation allows them to be treated as four separate 2×2 matrices. This allows us to make use of the results in Sec. II, which is computationally easier than solving the original eight-level system. To determine the trajectory sector the plaquette will lie in, we must first calculate Eq. (56) for all four sectors and use this as a weighting to randomly select one of the four sectors. Then, we can sample the edge configurations η_m after each time τ_m , as was done in Sec. IV. This gives a sequence of configurations $\{\eta_0, \dots, \eta_{\tilde{M}}\}$ which are sampled with probability

$$P(\eta_0, \cdots, \eta_{\tilde{M}}) = \frac{1}{Z_{\beta}^{\Delta}[\{\boldsymbol{\eta}^{\Delta'}\}]} \prod_{m=0}^{\tilde{M}} Z_{\Delta\tau_m}^{\Delta}(\eta'_m, \eta_{m+1}).$$
(57)

3. Mapping to a spin and sampling stochastic bridges

The partition function $Z^{\Delta}_{\Delta\tau_m}(\eta'_m, \eta_{m+1})$ describes the ensemble of all possible trajectories for the plaquette Δ between the times $\tau_m < t < \tau_{m+1}$. Within these times, the plaquette can only hop between two distinct configurations for some chosen η'_m . These configurations will have opposite magnetization (cf. Fig. 5). As such, it is convenient to treat the plaquette as a spin κ^{Δ} , with

$$\kappa^{\Delta}(t) = \operatorname{sgn}(\sigma^{i}(t) + \sigma^{j}(t) + \sigma^{k}(t)),$$
(58)

and an effective longitudinal field

$$g(t) = J|\sigma^{i}(t) + \sigma^{j}(t) + \sigma^{k}(t)|, \qquad (59)$$

which is constant between times $\tau_m < t < \tau_{m+1}$. We can now write

$$Z_{\Delta\tau_{m}}^{\Delta}(\eta_{m}',\eta_{m+1}) = \sum_{\{\eta_{m}^{\Delta}\}} \delta[\eta_{m}',\eta^{\Delta}(\tau_{m})] \,\delta[\eta_{m+1},\eta^{\Delta}(\tau_{m+1})]$$

$$\times e^{J\int_{\tau_{m}}^{\tau_{m+1}} dt \sum_{i \in \Delta} \sigma^{i}(t)} \prod_{k_{m}=1}^{\mathcal{K}(\eta_{k_{m}}^{\Delta})} K_{\eta_{k_{m}-1}^{\Delta} \to \eta_{k_{m}}^{\Delta}}$$

$$= \sum_{\{\kappa_{m}^{\Delta}\}} \delta[\kappa_{m}',\kappa^{\Delta}(\tau_{m})] \,\delta[\kappa_{m+1},\kappa^{\Delta}(\tau_{m+1})]$$

$$\times e^{g(\tau_{m})\int_{\tau_{m}}^{\tau_{m+1}} dt \,\kappa^{\Delta}(t)} \prod_{k_{m}=1}^{\mathcal{K}(\kappa_{k_{m}-1}^{\Delta} \to \kappa_{k_{m}}^{\Delta}}$$

$$:= Z_{\Delta\tau_{m}}^{\kappa}(\kappa_{m}',\kappa_{m+1};g(\tau_{m})), \qquad (60)$$

where $\eta_m^{\triangle} = \eta^{\triangle}(\tau_m : \tau_{m+1})$ and $\kappa_m^{\triangle} = \kappa^{\triangle}(\tau_m : \tau_{m+1})$ are partial trajectories between times τ_m and τ_{m+1} for the plaquette and spin, respectively.

Now that the partition function is expressed as the dynamics of a single spin, it can again be sampled using the results from Sec. III D with a longitudinal field $g(\tau_m)$, and then mapped back onto the plaquette. The probability of generating each stochastic bridge is

$$P(\boldsymbol{\eta}_{m}^{\bigtriangleup} \mid \boldsymbol{\sigma}^{(\bigtriangleup)}) = \frac{\delta[\boldsymbol{\eta}_{m}^{\prime}, \boldsymbol{\eta}(\tau_{m})] \,\delta[\boldsymbol{\eta}_{m+1}, \boldsymbol{\eta}(\tau_{m+1})]}{Z_{\bigtriangleup\tau_{m}}^{\bigtriangleup}(\boldsymbol{\eta}_{m}^{\prime}, \boldsymbol{\eta}_{m+1})} \\ \times e^{J \sum_{i \in \bigtriangleup} \int_{\tau_{m}}^{\tau_{m+1}} dt \, \sigma^{i}(t)} \prod_{k_{m}=1}^{\mathcal{K}(\boldsymbol{\eta}_{m}^{\bigtriangleup})} K_{\boldsymbol{\eta}_{k_{m}-1}^{\bigtriangleup} \to \boldsymbol{\eta}_{k_{m}}^{\bigtriangleup}}.$$
(61)

The probability of generating the full plaquette update is the product of sampling the edge configurations and then sampling stochastic bridges between them:

$$P(\boldsymbol{\eta}^{\triangle} \mid \boldsymbol{\sigma}^{(\triangle)}) = \frac{\delta[\boldsymbol{\eta}^{\triangle}(0), \, \boldsymbol{\eta}^{\triangle}(\boldsymbol{\beta})]}{Z_{\boldsymbol{\beta}}^{\triangle}[\{\boldsymbol{\eta}^{\triangle'}\}]} \times e^{J\sum_{i\in\Delta}\int_{0}^{\boldsymbol{\beta}} dt \, \sigma_{i}(t)} \prod_{k=1}^{\mathcal{K}(\boldsymbol{\eta}^{\triangle})} K_{\boldsymbol{\eta}_{k-1}^{\triangle} \rightarrow \boldsymbol{\eta}_{k}^{\triangle}}.$$
 (62)

4. Detailed balance

We can now verify that the single plaquette update obeys detailed balance and is thus rejection-free. Given some trajectory $\boldsymbol{\omega}$, the probability to generate some new trajectory $\tilde{\boldsymbol{\omega}}$ is

$$P(\tilde{\boldsymbol{\omega}} \mid \boldsymbol{\omega}) = N^{-1} P(\tilde{\boldsymbol{\eta}}^{\Delta} \mid \boldsymbol{\sigma}^{(\Delta)}), \tag{63}$$

where the factor N^{-1} comes from the fact the plaquette \triangle is chosen at random. Since $P(\eta^{\triangle} | \sigma^{(\triangle)}) \propto P_{\beta}(\omega)$, it follows that

$$\frac{P(\tilde{\boldsymbol{\omega}} \mid \boldsymbol{\omega})}{P(\boldsymbol{\omega} \mid \tilde{\boldsymbol{\omega}})} = \frac{P(\tilde{\boldsymbol{\eta}}^{\bigtriangleup} \mid \boldsymbol{\sigma}^{(\bigtriangleup)})}{P(\boldsymbol{\eta}^{\bigtriangleup} \mid \boldsymbol{\sigma}^{(\bigtriangleup)})} = \frac{P_{\beta}(\tilde{\boldsymbol{\omega}})}{P_{\beta}(\boldsymbol{\omega})},$$
(64)

and thus detailed balance is obeyed.

B. Thermal annealing

The expected first-order phase transition at J_c can slow down the convergence for $J \approx J_c$ and large inverse temperatures, β . In particular, if the initial seed trajectory is chosen to be in one of the two phases, there is the possibility that the update procedure cannot explore the entire trajectory space, thus remaining stuck in the incorrect phase. This is a consequence of the large barriers which need to be overcome to move between phases. Furthermore, making assumptions *a priori* on which phase the trajectory should belong to for some value of *J* could bias the results.

A common technique used in Monte Carlo sampling to overcome such metastability is *thermal annealing*. In this approach, we start from a small inverse temperature, $\beta = 0.1$, and gradually increase it to the target inverse temperature, $\beta = 128$. Our annealing schedule is to make N updates to the trajectory and then increase β by $\Delta\beta = 0.1$ for $\beta < 32$, and $\Delta\beta = 1$ for $\beta \ge 32$. When the inverse temperature is increased, we have to modify the trajectory to account for this. In practice, we just stretch the trajectory time by a factor of $(\beta + \delta\beta)/\beta$. After reaching the target inverse temperature, we do $N \times 10^2$ updates, and then restart the process. We repeat this procedure 10^3 times.

We find this approach to work well. Indeed, when sufficiently far from J = 1, the generated trajectories (at the target dynamics) have behavior corresponding to their correct phase with high accuracy. However, when close to J = 1, the trajectories can have properties which correspond to either phase (in practice, we find that the process picks just one of the phases for each run of the annealing process). While we cannot be certain this approach guarantees the correct amount of mixing between the phases, it provides a less biased way to propose initial trajectory seeds, and still demonstrates the first-order behavior of the transition point.

C. First-order quantum phase transition of the QTPM

We now demonstrate how the plaquette update scheme with temperature annealing can be used to investigate the quantum phase transition of the QTPM. As described in Ref. [24], the finite size study of the QTPM has to be done with care, since, depending on the size and aspect ratio of the system used, Eq. (50) can have many or no symmetries. We focus on square lattices $N = L \times L$, with $L = \{4, 5, 8, 11, 16, 23, 32\}$ chosen such that there are no such symmetries [24]. We also compare our results against those from 2D DMRG for $N = 4 \times 4$.

Figure 6(a) shows the average transverse three-spin correlator:

$$M_{xxx} = N^{-1} \sum_{\{i,j,k\} \in \Delta} \langle \hat{X}_i \hat{X}_j \hat{X}_k \rangle.$$
(65)

The inset shows the results close to the transition point. Figure 6(b) shows the average longitudinal magnetization:

$$M_z = N^{-1} \sum_{i=1}^N \langle \hat{Z}_i \rangle.$$
(66)

As we increase the system size, the crossover between the two phases becomes increasingly steplike, suggesting a first-order singularity in the large-size limit. The behavior of the magnetic susceptibility,

$$\chi_Z(J) = \frac{dM_z}{dJ},\tag{67}$$



FIG. 6. Ground-state phase transition in the QTPM. Results using TPS with exact plaquette updates and simulated annealing for $N = L \times L$ and $L = \{4, 5, 8, 11, 16, 23, 32\}$. We show (a) the three spin correlator, M_{xxx} ; (b) the longitudinal magnetization, M_z ; and (c) the magnetic susceptibility, $\chi_z(J) = dM_z/dJ$. The data points show results from trajectory sampling, and for comparison, the dashed lines show results from DMRG for $N = 4 \times 4$. The insets show the behavior close to the transition point.

is consistent with this interpretation, with its peak getting higher and narrower around J = 1 with the system size, see Fig. 6(c).

VI. CONCLUSIONS

In this paper, we have leveraged the connection between the continuous-time expansion of the quantum Boltzmann distribution and the rare-event sampling in stochastic dynamics for systems which have no sign problem. In particular, we have focused on the so-called stoquastic Hamiltonians, where computing the partition sum is equivalent to sampling imaginary-time trajectories of a continuous-time Markov chain. Each trajectory in the ensemble is conditioned to return to the initial configuration, and its probability is exponentially biased (or tilted) due to the difference between the Hamiltonian and the associated stochastic generator. Such trajectories can be accessed with a method like TPS, i.e., Monte Carlo for trajectory ensembles. Specifically, we showed that in systems with finite-state local degrees of freedom (such as spins on a lattice), one can use an approach similar to that of Refs. [17,18] to devise a rejection-free TPS scheme by means of an exact local generation of trajectory updates which is especially simple in spin-1/2 models. In fact, as we showed above, this gives the optimal, or Doob, dynamics for sampling the rare trajectories.

We illustrated the effectiveness of this approach by studying the quantum phase transitions of two classes of models. The first included the TFIM in 1D and 2D, where the transition is well known to be continuous. We showed that the rejection-free TPS method correctly characterises their quantum phase transition, even in the near-critical regime where trajectories take many TPS iterations to decorrelate. The second class of models we considered are quantum plaquette models. In particular, we studied the QTPM and showed how to generalize the rejection-free method to local multispin updates. While less understood than the TFIM, the OTPM has a quantum phase transition which is first-order. Again, the rejection-free method efficiently recovered the quantum phase transition. As a by-product, we also computed to high accuracy the statistics of dynamical observables (both in the LD regime and for finite imaginary times) in

the trajectory ensembles that resolve the quantum partition sums of both the TFIM and the QTPM, which is shown in Appendix B.

The method described here was used recently in Refs. [24,48], where the ground-state phase transitions of various spin models were characterized for large system sizes. We foresee that our approach will be useful for numerical investigations of glassy models, where nonlocal updates could be difficult to formulate. While the method we used here is based on local updates, the key property is not locality but simplicity, which allows the sampling of exact trajectory moves by solving Eqs. (21) and (22) that define the Doob dynamics. It would be interesting to find other (perhaps nonlocal) moves that also solve Eqs. (21) and (22). While finding exact updates might prove difficult, one might be able to formulate an approximately optimal dynamics, with the error accounted for in an acceptance test, such as Metropolis in a TPS scheme. The difficulty here is finding a dynamics which provides improvements in sampling while defined in terms of transition rates which remain computationally cheap (perhaps building on the use of tensor networks for optimal sampling [49]). A second proposition could be to implement more advanced sampling methods to overcome local minima, such as *parallel tempering*. While this could present some technical challenges, it might be advantageous for the sampling of glassy models.

The data shown in the figures is available at Ref. [50]. Example code used to generate the data can be found at Ref. [51].

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APPENDIX A: SAMPLING TIME-DEPENDENT OPTIMAL DYNAMICS

Suppose we are able to resolve the optimal dynamics given in Sec. III C, with the transition rates given by Eq. (21). This is a time-dependent dynamics which is defined through the reduced dynamics Eq. (15). To sample dynamics from it, we follow the process drawn out in Sec. III B, which requires randomly drawing waiting times from the distribution Eq. (17), and then randomly selecting a configuration to jump to from the distribution Eq. (18).

Determining a random transition time, τ , is easily done by considering the cumulative distribution function of Eq. (17):

$$C_x(\tau;\beta,x_{\rm f},t) = 1 - \exp\left[-\int_t^{t+\tau} dt' \tilde{R}_x(t';\beta,x_{\rm f})\right]. \quad (A1)$$

It is well understood that by drawing some uniformly random number $r \in [0, 1]$ and inverting $C_x(\tau; \beta, x_f, t) = r$, one can pick τ with a probability density given by Eq. (17). Using Eq. (22), we find

$$C_{x}(\tau;\beta,x_{\rm f},t) = 1 - \frac{Z_{\beta-t-\tau}(x;x_{\rm f})}{Z_{\beta-t}(x;x_{\rm f})} e^{-\tau H_{x}^{c}}.$$
 (A2)

While analytically inverting Eq. (A2) is difficult, this can be done numerically to arbitrary accuracy using the bisection method if one is able to calculate $Z_t(x; x_f)$. This is due to the fact that $C_x(\tau; \beta, x_f, t)$ is a monotonically increasing function between $C_x(0; \beta, x_f, t) = 0$ and $C_x(\beta - t; \beta, x_f, t) < 1$. Note that if $r > C_x(\beta - t; \beta, x_f, t)$, then no transition time is drawn, and the system remains in state x until the terminating trajectory time, β .

Once a transition time, τ , has been drawn, the state which the system transitions to can be determined using Eq. (18):

$$P_{x}(y \mid t + \tau; \beta, x_{f}) = \frac{K_{x \to y} Z_{\beta - t - \tau}(y; x_{f})}{\sum_{z \neq x} K_{x \to z} Z_{\beta - t - \tau}(z; x_{f})}.$$
 (A3)

APPENDIX B: TRAJECTORY STATISTICS

The connection between the quantum partition function and the ensemble of (conditioned/biased) stochastic trajectories naturally motivates the investigation of the *trajectory statistics* of the imaginary-time dynamics. Given some trajectory ensemble, such as one defined by Eq. (7), and some inverse temperature (or *trajectory time*), β , we can define its probability distribution function over a trajectory observable, $\mathcal{O}(\mathbf{x})$, through

$$P_{\beta}(\mathcal{O}) = \sum_{\{\mathbf{x}\}} \pi(\mathbf{x}) \delta[\mathcal{O}(\mathbf{x}) - \mathcal{O}].$$
(B1)

For simplicity, we will only consider trajectory observables which are obtained by time integrating diagonal operators:

$$\mathcal{O}(\mathbf{x}) = \int_0^\beta dt \ O(x(t)). \tag{B2}$$

In practice, the computation of Eq. (B1) for some arbitrary inverse temperature β is difficult. However, in the low-temperature limit, $\beta \rightarrow \infty$, one can estimate Eq. (B1) using the framework of LD theory (for reviews, see, e.g., Refs. [52–55]).

In the low-temperature limit, both the probability distribution of O and its moment generating function (MGF),

$$Z_{\beta}(s) = \sum_{\{\mathbf{x}\}} P_{\beta}(\mathbf{x}) e^{-s\mathcal{O}(\mathbf{x})}, \tag{B3}$$

take the LD forms

$$P_{\beta}(\mathcal{O}) \asymp e^{-\beta \varphi(\mathcal{O}/\beta)}, \quad Z_{\beta}(s) \asymp e^{\beta \theta(s)},$$
 (B4)

respectively. Here, φ is the *rate function*, and θ is the *scaled cumulant generating function*, with the two related through the Legendre transform, $\theta(s) = -\min_o[so + \varphi(o)]$, where $o = \mathcal{O}/\beta$.

The MGF has a form similar to the partition function, Eq. (3). In terms of trajectories, it reads cf. Eq. (4):

$$Z_{\beta}(s) = \sum_{\{x\}} P_{\beta}(x) e^{-s \int_{0}^{s} dt \, O(x(t))}$$

= $\frac{1}{Z_{\beta}} \sum_{\{x\}} \delta[x(0), x(\beta)] e^{-\int_{0}^{\beta} dt \, [H_{x(t)}^{c} + s \, O(x(t))]}$ (B5)

$$\prod_{m=1}^{\infty(\mathbf{x})} K_{x_{m-1} \to x_m}.$$
 (B6)

Furthermore, from Eq. (B6) we also see that the MGF can be written as

$$Z_{\beta}(s) = \frac{\operatorname{Tr}[e^{-\beta H_s}]}{Z_{\beta}},\tag{B7}$$

where \hat{H}_s is a *tilting* [52] of the original Hamiltonian:

$$\hat{H}_s = \hat{H} + s\hat{O}. \tag{B8}$$

This means that in the limit of small temperatures, Eq. (B4), we get that $\theta(s) = E_s - E$, where E_s is the ground-state energy of \hat{H}_s and E that of $\hat{H}_{s=0} = \hat{H}$. For the following, we use this to connect quantum phase transitions to transitions in the LD statistics of trajectory observables.

1. TFIM

We first use the methods of Sec. IV to investigate the trajectory statistics of the TFIM in 2D. The first trajectory observable we consider, cf. Eq. (B2), is the time integral of the two-point correlator:

$$\mathcal{M}_{zz}(\boldsymbol{\omega}) = \int_0^\beta d\tau \sum_{\langle i,j \rangle} \sigma^i(\tau) \, \sigma^j(\tau). \tag{B9}$$

We focus on the low-temperature limit, $\beta \gg 1$, corresponding to the ground-state behavior, where LD theory can be applied. The LD statistics are retrieved from the ground-state properties of the tilted Hamiltonian,

$$\hat{H}_{s} = -h \sum_{i=1}^{N} \hat{X}_{i} - (J-s) \sum_{\langle i,j \rangle} \hat{Z}_{i} \hat{Z}_{j},$$
 (B10)

which is the same as the original Hamiltonian with $J \rightarrow J - s$. Figure 7 shows the rate function, $-\varphi(\mathcal{M}_{zz}/\beta)$, cf. Eq. (B4), for the 2D TFIM with (a) $J/J_c = 1$, (b) $J/J_c = 0.5$, and (c) $J/J_c = 1.5$. At criticality, we observe a broadening in this distribution, demonstrating the divergence in correlation lengths. In contrast,



FIG. 7. Dynamical fluctuations of the imaginary-time trajectories of the 2D TFIM. The low-temperature trajectory statistics are encoded in the dynamical LDs. We show the rate function, $-\varphi(\mathcal{M}_{zz}/\beta)/N$, for the two-point correlator, \mathcal{M}_{zz} , for (a) $J/J_c = 1.0$, (b) $J/J_c = 0.5$, and (c) $J/J_c = 1.5$. The data points show results from trajectory sampling with $\beta = 128$. The dashed line shows the 2D DMRG for $N = 4 \times 4$. (d) Rate function $-\varphi(\mathcal{M}_{|z|}/\beta)/N$ for $J/J_c = 1.0$ and $\beta = 64$. The dotted line shows the rate function for the time integral of the longitudinal magnetization of a single spin from Eq. (23), with $\langle \hat{\sigma}^z \rangle = 0.45$.

away from the critical point the distributions become narrower.

The second trajectory quantity we consider is the timeintegral of the order parameter:

$$\mathcal{M}_{|z|}(\boldsymbol{\omega}) = \int_0^\beta d\tau \left| \sum_{i=1}^N \sigma^i(\tau) \right|.$$
(B11)

The corresponding tilted Hamiltonian is

$$\hat{H}_s = -h \sum_{i=1}^N \hat{X}_i - J \sum_{\langle i,j \rangle} \hat{Z}_i \hat{Z}_j + s \left| \sum_{i=1}^N \hat{Z}_i \right|.$$
(B12)

To simulate dynamics with Eq. (B12), we can use the single-spin update scheme previously described, but we must



FIG. 8. Large deviations in the long imaginary-time trajectories of the QTPM. The long imaginary-time trajectory statistics can be captured through the LDs. We show the rate function, $-\varphi(\mathcal{M}_z)/N$, for (a) J = 1.0, (b) J = 0.8, and (c) J = 1.2. The data points show results from trajectory sampling for $N = L \times L$ with L ={4, 5, 8, 11, 16, 23, 32}, and the dashed line from 2D DMRG for L = 4. For comparison, the dotted lines show the rate functions for the single spin Eq. (23) case, where g is chosen to fit the mean to the peak of the QTPM distributions.

now consider the trajectories of *all* other spins when constructing the time-dependent dynamics due to the coupling introduced via $M_{|z|}$ in Eq. (B12). While this makes the procedure more costly, we are still able to run dynamics for moderate *N*. We show the rate function in Fig. 7(d). Here, the broadening at criticality is more pronounced; for comparison, we show the rate function for the longitudinal magnetization of a single spin (dashed line). This broadening suggests a diverging magnetic susceptibility in the large size limit.

2. QTPM

We now investigate the effects of the first-order phase transition of the QTPM at the level of the imaginarytime trajectories. We consider the statistics of the trajectory observable:

$$\mathcal{M}_{z}(\boldsymbol{\omega}) = \int_{0}^{\beta} d\tau \sum_{i=1}^{N} \sigma^{i}(\tau).$$
 (B13)

They are encoded in the tilted Hamiltonian,

$$\hat{H}_{s} = -h \sum_{\{i,j,k\} \in \Delta} \hat{X}_{i} \hat{X}_{j} \hat{X}_{k} - (J-s) \sum_{i=1}^{N} \hat{Z}_{i}, \qquad (B14)$$

which corresponds to the original Hamiltonian with $J \rightarrow J - s$. Figure 8 shows the rate function $-\varphi(\mathcal{M}_z/\beta)$ for various J. The rate function at J_c is shown in Fig. 8(a): the rate function



FIG. 9. Trajectory statistics at finite imaginary times in the QTPM. The estimated distribution of the time-integrated magnetization for L = 16 for various finite times (inverse temperatures) at J = 1 from trajectory sampling. The curves are calculated using a Gaussian kernel with $\delta = 0.005$.

flattens with increasing system size, indicating the existence of large fluctuations. This behavior is characteristic of a (dynamical) first-order phase transition due to the coexistence of two distinct (dynamical) phases. For comparison, we show the LDs of the single-spin problem from Eq. (23) (dotted lines), where the value of g is chosen to fix the mean of the distribution. Figure 8(a) shows the LDs of the single-spin problem with $\langle M_z \rangle = 0.42$ and 0.8, which approximately match the tails of the distribution for the QTPM. The interpretation is clear: the two coexisting phases are homogeneous phases of distinct \mathcal{M}_z , and intraphase fluctuations are uninteresting (thus the modes are well approximated by a single spin); the broadening in φ in the QTPM is due to a Maxwell construct

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between these modes due to the fact that intermediate values of M_z are realized by coexistence, i.e., space-time regions of one phase separated by sharp interfaces from space-time regions of the other phase. In Figs. 8(b) and 8(c), we also show the rate functions away from the transition point for J = 0.8, 1.2, respectively. While there is a slight broadening in the tails of these distributions, they still describe single phases far from coexistence.

It is also possible to reasonably estimate the probability distribution $P_{\beta}(\mathcal{M}_z)$ for finite β through sampling. This is shown in Fig. 9 at the transition point for L = 16 for various inverse temperatures. We use our annealing strategy to sample $N_{\text{traj}} = 10^6$ trajectories for each inverse temperature. The results are used to approximate the probability distribution function using the Gaussian kernel,

$$P_{\beta}(\mathcal{M}_z) \sim \sum_{i=1}^{N_{\text{traj}}} \exp\left(-\frac{[\mathcal{M}_z - \mathcal{M}_z(\boldsymbol{\omega}^i)]^2}{2\delta^2}\right), \qquad (B15)$$

where $\mathcal{M}_z(\omega^i)$ is the time-integrated magnetization of the *i*th trajectory and $\delta = 0.005$ is the width of the filter. Notice that at small inverse temperatures, the distribution looks approximately Gaussian. However, with increasing β , the distribution becomes bimodal, explicitly demonstrating the coexistence of phases. For the largest inverse temperatures shown here, trajectories with phase coexistence (trajectories with time-averaged \mathcal{M}_z between the two phases) are improbable. The broadening observed in Fig. 8(a) can only be seen for inverse temperatures much larger than the ones considered here, where phase coexistence within trajectories can be realized by rare transitions between the two phases.

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