



# Single-species dinoflagellate cyst carbon isotope fractionation in core-top sediments: environmental controls, CO<sub>2</sub> dependency and proxy potential

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**Abstract.** Sedimentary bulk organic matter and various molecular organic components exhibit strong CO<sub>2</sub>-dependent carbon isotope fractionation relative to dissolved inorganic carbon sources. This fractionation ( $\epsilon_p$ ) has been employed as a proxy for paleo- $p\text{CO}_2$ . Yet, culture experiments indicate that CO<sub>2</sub>-dependent  $\epsilon_p$  is highly specific at genus and even species level, potentially hampering the use of bulk organic matter and non-species-specific organic compounds. In recent years, significant progress has been made towards a CO<sub>2</sub> proxy using controlled growth experiments with dinoflagellate species, also showing highly species-specific  $\epsilon_p$  values. These values were, however, based on motile specimens, and it remains unknown whether these relations also hold for the organic-walled resting cysts (dinocysts) produced by these dinoflagellate species in their natural environment. We here analyze dinocysts isolated from core tops from the Atlantic Ocean and Mediterranean Sea, representing several species (*Spiniferites elongatus*, *S.* (cf.) *ramosus*, *S. mirabilis*, *Operculodinium centrocarpum* sensu Wall and Dale (1966) (hereafter referred to as *O. centrocarpum*) and *Impagidinium aculeatum*) using laser ablation–nano-combustion–gas-chromatography–isotope ratio mass spectrometry (LA/nC/GC-IRMS). We find that the dinocysts produced in the natural environment are all appreciably more <sup>13</sup>C-depleted compared to the cultured motile dinoflagellate cells, implying higher overall  $\epsilon_p$  values, and, moreover, exhibit large isotope variability. Where several

species could be analyzed from a single location, we often record significant differences in isotopic variance and offsets in mean  $\delta^{13}\text{C}$  values between species, highlighting the importance of single-species carbon isotope analyses. The most geographically expanded dataset, based on *O. centrocarpum*, shows that  $\epsilon_p$  correlates significantly with various environmental parameters. Importantly, *O. centrocarpum* shows a CO<sub>2</sub>-dependent  $\epsilon_p$  above  $\sim 240 \mu\text{atm } p\text{CO}_2$ . Similar to other marine autotrophs, relative insensitivity at low  $p\text{CO}_2$  is in line with active carbon-concentrating mechanisms at low  $p\text{CO}_2$ , although we here cannot fully exclude that we partly underestimated  $\epsilon_p$  sensitivity at low  $p\text{CO}_2$  values due to the relatively sparse sampling in that range. Finally, we use the relation between  $\epsilon_p$  and  $p\text{CO}_2$  in *O. centrocarpum* to propose a first  $p\text{CO}_2$  proxy based on a single dinocyst species.

## 1 Introduction

Stable carbon isotope fractionation in marine autotrophs is governed for a large part by the carbon-fixing enzyme Ru-BisCO (e.g., Farquhar et al., 1989; Roeske and O'Leary, 1984), which implies most marine organic matter and therefore sedimentary marine organic matter is strongly <sup>13</sup>C-depleted with respect to the dissolved inorganic carbon (DIC) source (CO<sub>2</sub> (aq), HCO<sub>3</sub><sup>-</sup> or CO<sub>3</sub><sup>2-</sup>), with the stable carbon isotope composition ( $\delta^{13}\text{C}$ ) of organic matter ranging from

–10‰ to –30‰ (Freeman and Hayes, 1992). While many groups of marine autotrophs show clear CO<sub>2</sub>-dependent carbon isotope fractionation ( $\epsilon_p$ ), the exact relation strongly varies between marine phytoplankton groups, genera and cell morphologies (Popp et al., 1998; Boller et al., 2011, 2015; Brandenburg et al., 2022). Still, because of the assumed CO<sub>2</sub> dependency of RuBisCO fractionation, bulk marine organic matter and more specific organic compounds of marine autotrophs (e.g., lipids biomarkers) have been proposed and applied as  $p$ CO<sub>2</sub> proxies over the past decades (Freeman and Hayes, 1992; Naafs et al., 2016). The application of these  $p$ CO<sub>2</sub> proxies (e.g., Bijl et al., 2010; Pagani et al., 2011; Schoon et al., 2011; Witkowski et al., 2018) has provided constraints on past atmospheric  $p$ CO<sub>2</sub> and earth system sensitivity beyond the ice core record (e.g., Pagani et al., 2010; PALAEOSSENS, 2012).

However, many of the organic compounds used for CO<sub>2</sub> reconstructions such as alkenones (e.g., Pagani, 2013), phytane (e.g., Witkowski et al., 2018), porphyrins (e.g., Freeman and Hayes, 1992) or bulk organic matter (e.g., Hayes et al., 1999) are not unique to a single species, genus and sometimes not even a group of organisms. This implies that reconstructions based on these compounds integrate interspecific differences in CO<sub>2</sub> dependency, which complicates the interpretation of such proxy records. Secondly, even if specific compounds derive from a single species or genus, they intrinsically derive from a multitude of individual organisms, differing in shape and size, affecting isotopic fractionation and hence limiting the accuracy of such CO<sub>2</sub> reconstructions.

Part of the uncertainties and biases in carbon isotope fractionation can be circumvented if the carbon isotopic fractionation of individual fossils can be analyzed. In recent years, significant progress has been made towards a CO<sub>2</sub> proxy based on the stable carbon isotope fractionation in organic walled dinoflagellate cysts (Burkhardt et al., 1999; Hoins et al., 2015, 2016a, b; Wilkes et al., 2017). A fraction (~15%) of modern dinoflagellates produces an organic resting cyst or dinocyst as an obligatory part of their lifecycle (Evvitt, 1985). The organic resting cysts from autotrophic species have excellent preservation potential and are often highly oxidation-resistant (Zonneveld et al., 1997, 2019; Kodrans-Nsiah et al., 2008), and several ubiquitous extant genera and species, such as *Spiniferites* spp. and *Operculodinium centrocarpum*, have extremely long geological records (Fensome et al., 1996; Williams et al., 2004). The ecology and morphology of these long-ranging species seemingly remained unchanged for millions of years (Frieling and Sluijs, 2018). Most importantly, recent advances in methodology allow for analyses of species-specific single-cyst  $\delta^{13}\text{C}$  (van Roij et al., 2017; Sluijs et al., 2018). This presents the opportunity to quantify environmental controls on  $\epsilon_p$  of individual dinoflagellate cysts and hence species to assess the potential to obtain more accurate paleo- $p$ CO<sub>2</sub> estimates from sedimentary records.

Controlled growth experiments across a range of CO<sub>2</sub> levels representative of the last glacial (e.g., Barnola et al., 1987) and modern and future carbon emission scenarios (Eberlein et al., 2016; Hoins et al., 2016a, b, 2015; IPCC, 2014; Rost et al., 2006; Van de Waal et al., 2013; Wilkes et al., 2017) showed species-specific CO<sub>2</sub>-dependent  $\epsilon_p$  for multiple dinoflagellate species. From these, the species *Protoceratium reticulatum* and *Gonyaulax spinifera* (Hoins et al., 2015, 2016a, 2016b) are of particular interest as these produce the organic cyst species *Operculodinium centrocarpum* sensu Wall and Dale (1966), hereafter referred to as *O. centrocarpum*, and *Spiniferites* (cf.) *ramosus*, hereafter referred to as *S. ramosus*, respectively (Head, 1996; Zonneveld et al., 2013). These cyst species have their first occurrences in the geological records around ~60 and 130 Ma (million years ago), for *O. centrocarpum* and *S. ramosus*, respectively (Williams et al., 2004), thus providing the potential for deep-time  $p$ CO<sub>2</sub> reconstructions.

Before  $\epsilon_p$  values based on dinocysts can be used for reconstructing  $p$ CO<sub>2</sub>, several fundamental questions need to be addressed. Although  $\delta^{13}\text{C}_{\text{DIC}}$  exerts a major control on dinocyst  $\delta^{13}\text{C}$  (Sluijs et al., 2018), it remains uncertain whether the CO<sub>2</sub> control on  $\epsilon_p$  of motile cells from controlled growth experiments can be translated to their cysts formed in the natural environment. In addition, potential offsets in  $\epsilon_p$  values between the motile cells and the cysts need to be established. This is especially important as the cell–cyst relations in carbon isotope ratios are not necessarily straightforward because bulk biomass such as cysts potentially deviates in  $\delta^{13}\text{C}$  values from the various cell components and potentially not by a constant offset (e.g., Freeman, 2001; Hayes, 2001; Pancost and Pagani, 2006; Schouten et al., 1998; Van de Waal et al., 2013; Wilkes et al., 2018).

We here present the first core-top data for single-species dinocysts to constrain the environmental controls on  $\epsilon_p$ . We focus on the species *O. centrocarpum* and compare this data, when possible, to several species of *Spiniferites* (*S. ramosus*, *S. mirabilis*, *S. elongatus*) and *Impagidinium aculeatum*. The established environmental relations are subsequently evaluated using simple models converting carbon isotope fractionation in dinocysts into  $p$ CO<sub>2</sub> values for the surface waters.

## 2 Materials and methods

### 2.1 Materials

The primary dataset is based on 34 core-top samples (Table 1, Fig. 1), collected from the North Atlantic Ocean and Mediterranean Sea, with a secondary dataset based on pre-industrial sample material spanning ~0–1500 common era (CE) from ENAM9606 for the North Atlantic (Richter et al., 2009). The core-top samples encompass a substantial natural  $p$ CO<sub>2</sub> (aq) gradient because the rate of cooling of the North Atlantic Current exceeds that of CO<sub>2</sub> uptake, whereas  $p$ CO<sub>2</sub>

in the Mediterranean is close to or slightly above equilibrium with the atmosphere. Sample selection is further based on the dinocyst occurrence maps of Zonneveld et al. (2013), including only samples with an expected relative abundance of at least 10 %–20 % of the target species. Similarly, the coverage of environmental parameters such as sea surface temperature (SST),  $p\text{CO}_2$  and difference in environmental settings was maximized during sample selection. Existing ocean databases are used for obtaining the relevant environmental parameters (Table 1).

## 2.2 Methods

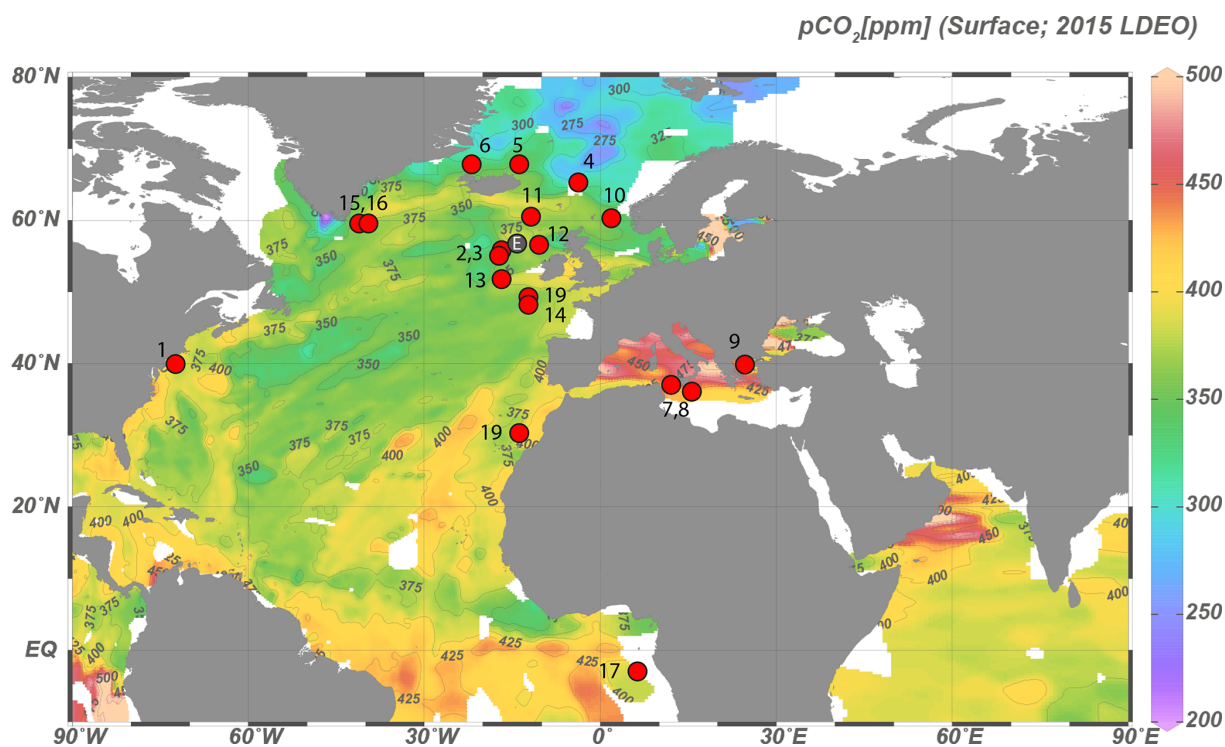
Using standard palynological techniques (see, e.g., Brinkhuis et al., 2003), ca. 5–10 g freeze-dried sediment of the upper 1–2 cm of core material was processed for each sample. This involved dissolving carbonates and silicate components using strong acids (HCl 30 % and HF 38 %–40 %). After acid steps, residues were pH-neutralized and sieved using an ultrasonic bath and 250 and 10  $\mu\text{m}$  nylon mesh sieves to remove large and small particles, respectively. Subsequently, samples were transferred to glass test tubes with demineralized water and centrifuged at 3200 rpm for 10 min to obtain an optimum concentrate of the sample material. Prior to dinocyst selection, samples were stored in 4 mL glass vials in demineralized water.

A micro-manipulator consisting of a Leica inverted microscope and a Narishige IM-9B micro-injector connected to a strung-out pipette was used to manually select individual dinocysts from a water droplet on a glass petri dish. Dinocyst selection followed a strict protocol, in which cyst morphology (primarily cyst shape and process length) was kept constant and the contribution of other organic particles was minimized. Specimens with darker coloration or amorphous organic matter adhering to the cyst or processes were avoided. In the case of *O. centrocarpum*, the morphological selection primarily involved selecting specimens of equal size and process length to avoid cysts that may be derived from different environments (e.g., Mertens et al., 2009). For *Spiniferites*, we were able to distinguish and separate three distinct morphological species in sufficient numbers: *S. ramosus*, *S. elongatus* and *S. mirabilis*. For all dinocyst species, the selected diameter excluding processes was in the order of  $\sim 30$ – $40 \mu\text{m}$ , except for *S. mirabilis* ( $\sim 60 \mu\text{m}$ ), although constraining the exact size of each individual specimen was not feasible within the current analytical procedures. Stable carbon isotope analyses for individual samples are based on replicating the analyses of single of dinocysts, with  $\sim 30$  individual measurements being conducted to obtain a reasonably precise ( $\sim 0.3 \text{‰}$ – $0.4 \text{‰}$ ) sample average (van Roij et al., 2017). Given the size of the dinocysts used here ( $\sim 30$ – $40 \mu\text{m}$  cyst diameter), three to seven specimens were required for each measurement, and hence  $\sim 150$  cysts were required to obtain sample averages (Table 1).

Dinocysts were placed on a 6 mm  $\varnothing$  nickel sample tray, after which an identical second tray is added on top and compressed to fixate the dinocysts. Before placement in the ablation chamber, approximately  $\sim 1 \text{mm}^2$  of International Atomic Energy Agency CH-7 (IAEA-CH7) polyethylene standard (PE; certified  $\delta^{13}\text{C}$  value  $-32.15 \text{‰} \pm 0.05 \text{‰}$ ;  $1\sigma$ ) was added to the sample tray. Stable carbon isotope analyses of the dinocysts followed the procedures described in previous work (van Roij et al., 2017; Sluijs et al., 2018), utilizing the recently developed ablation–nano-combustion–gas-chromatography–isotope ratio mass spectrometry (LA/nC/GC-IRMS) method. Fragments resulting from deep-ultraviolet LA were carried using a continuous helium flow in 0.32 mm capillaries and oxidized in a combustion oven at  $940^\circ\text{C}$ . The resultant  $\text{CO}_2$  was transported to a GC combustion interface, dried in a Nafion tube using a He counterflow and subsequently into a Thermo Fisher DeltaV Advantage IRMS for isotope analysis. Each analytical run included five standards with signal intensity above 4 Vs (ca. 40 ng C;  $\delta^{13}\text{C}$  precision better than  $0.5 \text{‰}$ ) to allow calibrating to the Vienna Pee Dee Belemnite (VPDB) scale. Direct visual monitoring of the ablation process was used as initial quality assessment of each individual measurement.

To calculate the fractionation factor  $\varepsilon_p$  of the dinocysts relative to dissolved inorganic carbon (DIC) from which the dinocyst was produced, we take the  $\delta^{13}\text{C}_{\text{DIC}}$  from the modeled grid of Tagliabue and Bopp (2008). As many dinoflagellate species, including those that produce *O. centrocarpum* and *S. ramosus* cysts, are able to utilize both  $\text{HCO}_3^-$ , which makes up the majority of DIC, and  $\text{CO}_2$  for carbon fixation (Hoins et al., 2016a), we also compare the  $\delta^{13}\text{C}_{\text{DINO}}$  data to  $\delta^{13}\text{C}_{\text{CO}_2}$  and to overall seawater carbon partitioning.

$\varepsilon_{p\text{-DIC}}$  is calculated as  $\delta^{13}\text{C}_{\text{DIC}} - \delta^{13}\text{C}_{\text{DINO}}$ , and  $\varepsilon_{p\text{-CO}_2}$  is calculated as  $\delta^{13}\text{C}_{\text{CO}_2} - \delta^{13}\text{C}_{\text{DINO}}$ . For the latter the  $\delta^{13}\text{C}$  of dissolved  $\text{CO}_2$  is calculated from  $\delta^{13}\text{C}_{\text{DIC}}$  using the temperature-dependent fractionation between DIC and  $\text{CO}_2(\text{aq})$  (Mook et al., 1974). To evaluate the dominant contributions to  $^{13}\text{C}$  fractionation in dinocysts, we compare the  $\varepsilon_{p\text{-DIC}}$  and  $\varepsilon_{p\text{-CO}_2}$  values to measured and interpolated physicochemical parameters. We test a suite of parameters,  $[\text{NO}_3^-]$ ,  $[\text{PO}_4^{3-}]$ ,  $[\text{Si}]$ , alkalinity,  $p\text{CO}_2$ , SST and sea surface salinity (SSS), which are extracted using Ocean Data View (<https://odv.awi.de/>, last access: 20 July 2020) from existing (gridded) datasets (Gouretski and Koltermann, 2004; Takahashi et al., 2014, 2016) (Supplement). Where possible, data are averaged over a grid of  $4^\circ$  longitude and latitude around the sample position. This is both to reduce errors introduced by data scarcity and to account for potential lateral transport of dinocysts during sinking (Nooteboom et al., 2019). Carbonate chemistry is calculated using the R package seacarb (Gattuso et al., 2021), with alkalinity and  $p\text{CO}_2$  as input variables to calculate the other relevant parameters of the carbonate system: the relative contributions of  $\text{CO}_2(\text{aq})$ ,  $\text{HCO}_3^-$  and  $\text{CO}_3^{2-}$ , i.e., carbon speciation.



**Figure 1.** Locations of samples with sufficient *Operculodinium centrocarpum* and/or *Spiniferites* spp for dinoflagellate cyst  $\delta^{13}\text{C}$  analyses. Numbers correspond to localities listed in Table 1. Down-core data location (ENAM9606;  $55.650^\circ\text{N}$ ,  $13.985^\circ\text{W}$ ) is marked by a gray dot (“E”) between core top nos. 2, 3 and 12.

Ideally, all environmental parameters would be known for the different locations as would the time the dinoflagellates lived and encysted. This is, however, unfeasible because the dinocysts assemblage in core-top sediments (typically the upper 2 cm of sediment) integrates conditions over several centuries, assuming moderate to low average sediment accumulation rates ( $< 10\text{ cm kyr}^{-1}$ ) that generally characterize open-ocean settings such as examined here. We therefore apply a rough correction for  $p\text{CO}_2$ , based on the assumption that local air–sea gas exchange has remained similar, that equals the atmospheric  $p\text{CO}_2$  rise between the sampling date and 1850 CE. The correction from actual measurements to “pre-industrial” conditions for  $p\text{CO}_2$  yields a substantial offset due to the  $\sim 90\text{ ppmv}$  atmospheric  $p\text{CO}_2$  rise from 1850 CE to the average sampling date (ca. 2000 CE). As this correction is broadly similar for all sample localities, it has only a small impact on the overall pattern in the  $\text{CO}_2$  data (Fig. 3). We employed a Monte Carlo simulation to assess the potential impact of the  $p\text{CO}_2$  correction by propagating (1) the 5% analytical error in  $p\text{CO}_2$  values and adding  $45 \pm 15\text{ ppm}$  to reflect a normally distributed mixture of modern and pre-industrial conditions and (2) a resampled uncertainty derived from the  $p\text{CO}_2$  rise since pre-industrial times (1800–2000 CE). Both scenarios are set up to represent worst-case scenarios; a single error drawn from the error distribution is imposed on a sample basis and not a resampled

average of the number of  $\delta^{13}\text{C}_{\text{DINO}}$  measurements within a sample, as that would reduce the error through averaging.

Changes in SST, SSS and nutrient concentrations are also expected, partly also by anthropogenic activity, but offsets in these parameters are generally subtle and more local compared to the changes in  $p\text{CO}_2$  and hence would require site-specific reconstructions. Still, recent widespread eutrophication and enhanced productivity may impact the carbon isotope results through increased DIC uptake in algal blooms (i.e., counteracting the impact of enhanced  $p\text{CO}_2$ ). However, as eutrophication has mainly affected coastal areas (Hallegraeff, 1993; Anderson et al., 2002), this is expected to play a minor factor at our, mostly open marine, sample localities (Fig. 1). Lastly, long-term natural changes in nutrients, SSS and SST also occur, and it is currently not possible to fully filter out the various anthropogenic offsets. With the exception of  $p\text{CO}_2$ , we hence assume all parameters (SST, SSS, nutrients) to have remained constant over the period the core-top samples represent.

### 3 Results

#### 3.1 Carbon yields from dinocyst analyses

Despite our pre-screening to include only samples with high relative abundances of the target species, some of the se-

**Table 1.** Core localities, analyzed species, number of measurements and normality of the carbon isotope data. Number of measurements is given as the total number and, in parentheses, the number of measurements used for environmental comparisons (see also results Sect. 3.4). Shapiro–Wilk (S–W) normality test on data: non-normal data distributions are indicated where  $p$  values are \* < 0.1, \*\* < 0.01 and \*\*\* < 0.001; parentheses indicate the same for the data used for environmental comparisons. Site numbers correspond to those in Figs. 1 and 6a.

Site number	Core ID	Latitude	Longitude	Species	Measurements ( $n$ )	S–W normality
1	NF2012-091	37.977402° N	73.669403° W	<i>O. centrocarpum</i>	22 (21)	*
2	PE360-24	55.496231° N	15.800755° W	<i>O. centrocarpum</i>	24 (23)	**
3	PE360-45	55.539398° N	15.8453° W	<i>O. centrocarpum</i>	23 (16)	
4	NA87-02	64.480003° N	5.83° W	<i>O. centrocarpum</i>	20 (14)	***
5	LCD13	67.501282° N	15.069252° W	<i>O. centrocarpum</i>	25 (25)	
6	LCD10A	66.677437° N	24.179598° W	<i>O. centrocarpum</i>	29 (27)	
7	MedSea (MC-613)	35.8575° N	14.105556° E	<i>O. centrocarpum</i>	35 (35)	
8	MedSea (MC-614)	35.8075° N	12.998056° E	<i>O. centrocarpum</i>	33 (32)	*
9	MedSea (MC-645)	40.2175° N	25.244167° E	<i>O. centrocarpum</i>	33 (23)	**
10	ENAM93-08bx	59.501667° N	3.69° E	<i>O. centrocarpum</i> <i>S. ramosus</i>	38 (37) 33 (32)	*
11	ENAM94-13bx	60.249997° N	11.19° W	<i>O. centrocarpum</i> <i>S. elongatus</i> <i>S. mirabilis</i>	36 (33) 34 (30) 13 (12)	***
12	ENAM96-09bx	57.159917° N	10.26° W	<i>O. centrocarpum</i> <i>S. ramosus</i>	39 (37) 30 (29)	
13	ENAM97-04	52.410386° N	14.94° W	<i>O. centrocarpum</i> <i>S. elongatus</i> <i>S. mirabilis</i>	52 (50) 23 (23) 38 (37)	** (*)
14	Omex93-A2 bx	49.483° N	11.13° W	<i>O. centrocarpum</i> <i>S. ramosus</i> <i>S. elongatus</i> <i>S. mirabilis</i>	30 (29) 13 (11) 6 (6) 18 (17)	*
15	PE275-6	59.272369° N	38.36° W	<i>O. centrocarpum</i> <i>S. ramosus</i>	34 (33) 34 (30)	* **
16	PE275-9	59.272369° N	38.36° W	<i>O. centrocarpum</i> <i>S. ramosus</i> <i>S. elongatus</i>	37 (36) 23 (22) 12 (10)	* (*) * (*) * (*)
17	T89-15bx	4.199372° S	10.05° E	<i>O. centrocarpum</i> <i>S. ramosus</i>	36 (36) 35 (34)	* (*) * (*)
18	64PE428-1-1-6	47.079782° N	10.197305° W	<i>O. centrocarpum</i> <i>I. aculeatum</i>	35 (33) 30 (29)	
19	64PE428-1-6-6	30.67917° N	11.930478° W	<i>O. centrocarpum</i>	11 (10)	

lected samples contained too few dinocysts or at too low an abundance relative to other organic particles (amorphous organic matter, plant debris, pollen, non-dinocyst marine palynomorphs etc.) to be used in our study. Ultimately, out of the initial core-top sample set of 34 samples, 19 were found

suitable for species-specific dinocyst stable carbon isotope analyses (Table 1).

Typically, ~150 individual cysts were picked and analyzed for a total of 20–50 measurements, amounting to three to seven cysts per carbon isotope measurement. We calculate an average signal size of 0.2 Vs for a single cyst,

which amounts to a carbon yield of  $\sim 6\text{--}7$  ng C per cyst (van Roij et al., 2017). Although the variability in signal intensity from individual measurements suggests there is substantial intra-sample (cyst–cyst) variability, no significant off-sets in average carbon content per cyst were observed between samples, suggesting the average carbon content of the cysts within each of the analyzed populations is similar. *Spiniferites mirabilis* is the notable exception to this rule, as far fewer specimens of this species are needed for a single  $\delta^{13}\text{C}_{\text{cyst}}$  measurement. Based on the signal intensity per specimen, we estimate that this larger cyst species contains twice the amount of C compared to *S. ramosus*, *S. elongatus* and *O. centrocarpum*.

## 3.2 Carbon isotope data

### 3.2.1 Signal intensity

The 949 individual core-top analyses range in  $\delta^{13}\text{C}$  from  $\sim -18.5\text{‰}$  to  $-35.5\text{‰}$  (all  $\delta^{13}\text{C}$  are relative to Vienna Pee Dee Belemnite (VPDB)), and the 137 down-core analyses range from  $\sim -19.4\text{‰}$  to  $-28.1\text{‰}$ . No relation was observed between  $\delta^{13}\text{C}$  and signal size (Vs), except at the very low end ( $\leq 0.2$  Vs) (Fig. 2), in line with earlier analyses (van Roij et al., 2017). In this low range, the median of the  $\delta^{13}\text{C}$  values rises from  $-28\text{‰}$  below 0.1 Vs to values between  $-22\text{‰}$  and  $-25\text{‰}$  above 0.2 Vs. In the  $\leq 0.2$  Vs range the  $\delta^{13}\text{C}$  averages of both the cysts and PE converge between  $-25\text{‰}$  and  $-30\text{‰}$ , with substantial scatter. Poorer performance at such low C masses and signal intensities is expected, as these extremely small signal sizes and poor signal-to-noise ratio (below  $\sim 3:1$ ) approach the limit of our method. Consequently, even a very minor contamination source would bias values and result in larger scatter, as is also apparent in the PE standard at a similar signal intensity (van Roij et al., 2017; Fig. 2). Due to a worsening signal-to-noise ratio, we find that a noticeable degree of  $\delta^{13}\text{C}$  biasing from a background C source within the system is likely to occur at signal intensities  $\leq 0.5$  Vs and is particularly pronounced  $\leq 0.2$  Vs, and values for the standard and dinocysts converge around  $-27\text{‰}$  in this range (Fig. 2a). A similar background  $\delta^{13}\text{C}$  value was also obtained after liquid nitrogen trapping (van Roij et al., 2017). The source of this C remains elusive. It is unlikely to be related to the ablation (etching) of the nickel plate or associated with the water used to pick samples from, as measured blanks for those result in much lower signal intensities and neither source would affect the measurements of the PE standard. Lastly, a significant contribution of atmospheric  $\text{CO}_2$  ( $\delta^{13}\text{C}$  around  $-8\text{‰}$ ) appears unlikely due to the  $\delta^{13}\text{C}$  signature of the background signal ( $-27\text{‰}$ ). Though the origin of the background C remains unknown, we can use the trapping experiment to estimate the relevant background contribution (van Roij et al., 2017). We calculate the typical contribution is likely between 0.024 and 0.08 Vs, given a background C flux of 0.0008 Vs per second (van Roij

et al., 2017) and a typical duration of measurements (30–50 s for  $\delta^{13}\text{C}_{\text{DINO}}$  and up to  $\sim 100$  s for PE standard).

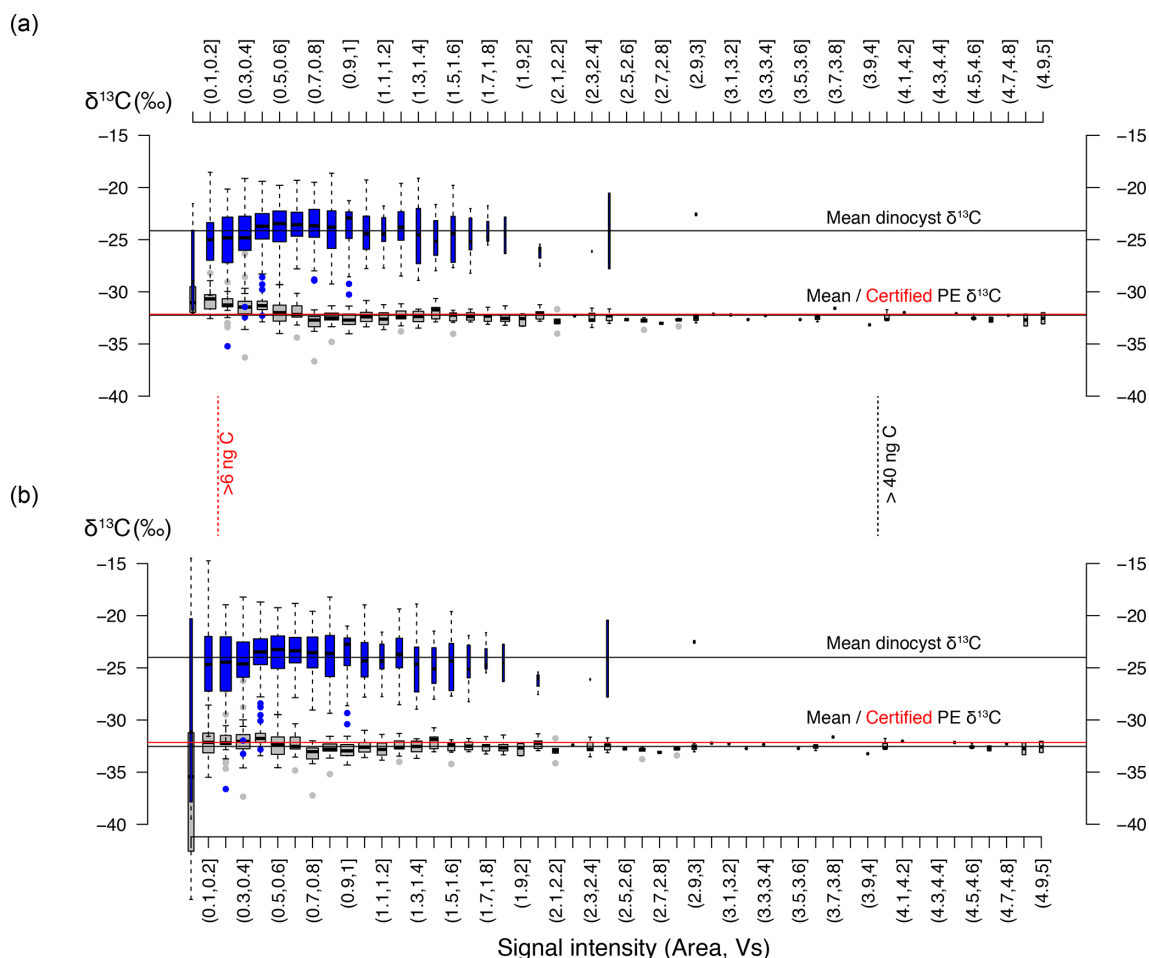
Before comparing our data with environmental variables, we therefore assess the impact of a very minor, but consistent, background contamination on the carbon isotope signal at low signal intensities (e.g., Fig. 2a). We find that a constant addition of ca. 0.04 Vs ( $\leq 1$  ng C) of a background C source with a  $\delta^{13}\text{C}$  of  $-27\text{‰}$  can explain the positive skewing in the standard PE  $\delta^{13}\text{C}$ . Using a simple isotope endmember and mass balance mixing model to correct for skewing (Fig. 2b), we calculated an average deviation from the measured PE and dinocyst values for intensities below 0.2 Vs in the order of  $|2.6\text{‰}|$  and  $|1.3\text{‰}|$ , respectively. The standard deviation of the data increases approximately 3-fold (Fig. 2b) compared to the raw measurement data below 0.2 Vs but remains virtually unchanged above 0.2 Vs, and the calculated deviation from the measured value is also much reduced above 0.2 Vs ( $< 0.3\text{‰}$ ).

The data correction using our simple mixing model eliminates the skew towards  $-27\text{‰}$  at low signal intensities and removes signal size  $\delta^{13}\text{C}$  dependency below 0.2 Vs for both the isotopically homogeneous PE and the heterogeneous dinocyst data (Fig. 2a, b). This suggests that our method of bias correction is warranted, but the increased variability at very low intensities and lack of independent control on the exact size and  $\delta^{13}\text{C}$  of the background contamination imply that the data associated with the lowest signal intensities remain significantly less reliable. We therefore apply a conservative cut-off and use only corrected data with a signal size above 0.2 Vs.

### 3.2.2 Skewed distributions and outlier omission

The drift-corrected  $\delta^{13}\text{C}_{\text{DINO}}$  is non-normally distributed in many core-top samples and also in different species (Table 1, Fig. 4). Distributions tend to be tailed towards lower values, exaggerated by the presence of a small number of outlier values. This is not due to analytical error or otherwise directly related to low signal intensity as we used a 0.2 Vs cut-off to eliminate samples with potentially unreliable signal-to-noise ratios (see above), and a minor correction for background C addition was sufficient to eliminate skewing at low signal intensities. The absence of such signals in the down-core samples (Supplement Fig. S1) suggests that the outliers and skewing in the sampled core-top populations could represent a real signal.

Based on typical deep-ocean sedimentation rates in the range of centimeters per kiloyear, the core-top samples are expected to contain a mixed assemblage of dinocysts produced mostly within the last centuries to millennia but could also include cysts produced during the last few decades that are likely affected by anthropogenic influences. It is particularly relevant to consider because a steep  $\delta^{13}\text{C}$  decrease ( $\sim 2\text{‰}$  since 1850 CE, of which  $> 1.5\text{‰}$  occurs after 1950 CE) (Francey et al., 1999; Keeling et al., 2017) accom-

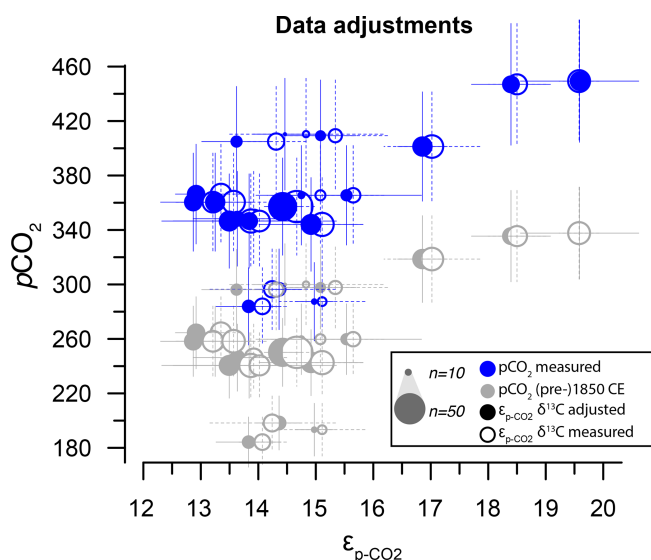


**Figure 2.** Signal intensity in volt seconds (Vs) versus carbon isotope value distribution within bins of 0.1 Vs. **(a)** The blue boxplots represent our new dinocyst  $\delta^{13}\text{C}$ , and the gray boxplots represent the previously published data for the PE standard ( $\delta^{13}\text{C}$  value  $-32.15\text{‰} \pm 0.05\text{‰}$ ;  $1\sigma$ ). **(b)** Same as **(a)**, but for background-corrected values (see Sect. 3.2). The vertical black and red lines represent cut-off values for individual PE standard measurements ( $>4 \text{ Vs}$ ;  $\sim 42 \text{ ng C}$ ;  $0.5\text{‰}$  precision) and individual cyst measurements ( $>0.2 \text{ Vs}$ ;  $\sim 6 \text{ ng C}$ ), respectively. The width of each boxplot is square-root-scaled with the number of measurements in the respective bins. Note that several bins at the high end do not contain any data.

panies the  $p\text{CO}_2$  rise ( $> 130 \text{ ppmv}$  since 1850 CE, of which  $> 100 \text{ ppmv}$  after 1950 CE). So even if enhanced carbon isotope fractionation at higher  $p\text{CO}_2$  (Freeman and Hayes, 1992; Hoins et al., 2015; Brandenburg et al., 2022) did not play a role, the most recent specimens are likely to be impacted by decreasing  $\delta^{13}\text{C}_{\text{DIC}}$ .

As age integration in the modern era may result in a mixture of cysts representing a range of environmental conditions, especially with respect to  $\text{CO}_2$  concentrations and  $\delta^{13}\text{C}_{\text{DIC}}$ , it is important to consider the potential age distribution of dinocysts before comparing  $\delta^{13}\text{C}_{\text{DINO}}$  and  $\varepsilon_p$  to environmental variables. In an ideal scenario, cysts produced after 1850 CE should be avoided in proxy-calibration efforts to eliminate a systematic bias towards the most recent times when atmospheric  $\text{CO}_2$  was already elevated above pre-industrial Holocene background ( $\sim 280 \text{ ppmv}$ ). Because

an accurate age correction for the Suess effect is technically unfeasible because the age distribution of  $\delta^{13}\text{C}_{\text{DINO}}$  measurements cannot be constrained, we illustrate the influence of  $\delta^{13}\text{C}_{\text{DINO}}$  data treatment (Fig. 3) and use Monte Carlo simulations of different error distributions to test the potential impact of the  $p\text{CO}_2$  correction. We also compared both measured  $p\text{CO}_2$  and  $p\text{CO}_2$  around 1850 CE (see Sect. 2.2) to  $\varepsilon_p$  calculated using both our raw  $\delta^{13}\text{C}_{\text{DINO}}$  data and the  $\delta^{13}\text{C}_{\text{DINO}}$  data after drift correction and removal of statistical outliers identified within the sample-specific single-species populations. This final step of data treatment removed positive and negative measurement outliers from the sample- and species-specific  $\delta^{13}\text{C}$  population (outside  $\pm 2.5$  IQR), after eliminating the extremely low-signal intensities ( $< 0.2 \text{ Vs}$ ) and correcting for the drift induced by background C in the system.



**Figure 3.** Effects of  $\delta^{13}\text{C}$  and  $p\text{CO}_2$  corrections (see also Fig. 2 and Sect. 3.2).

Altogether, out of a 949 core-top measurements, we omit 43 measurements with signals  $< 0.2$  Vs and 24 statistical outliers, which leaves 882 individual  $\delta^{13}\text{C}_{\text{DINO}}$  measurements: 560 for *O. centrocarpum*, 293 for *Spiniferites* (158 *S. ramosus*, 69 *S. elongatus* and 66 *S. mirabilis*) and 29 for *I. aculeatum* (Table 1). Most of the 67 omitted measurements have comparatively low  $\delta^{13}\text{C}$ , and the resulting  $\delta^{13}\text{C}$  of the populations is close to statistically indistinguishable from a normal distribution (Shapiro–Wilk  $p = 0.05\text{--}0.1$ ) or representative of a normal distribution (Table 1). Although the data treatment partly removed the negative skew on the  $\delta^{13}\text{C}_{\text{DINO}}$  distribution (Table 1), the combined effects of drift correction and outlier removal on sample and species-mean  $\delta^{13}\text{C}_{\text{DINO}}$  are generally small (Fig. 3). This is expected as drift correction averages only  $\sim 0.25\%$ , and the negative and positive outliers represent only a small percentage of the total measurements (Table 1).

Distinctly non-normally distributed  $\delta^{13}\text{C}$  values were not previously observed in recent pollen and ancient dinocyst species analyzed with the same method (van Roij et al., 2016; Sluijs et al., 2018). The here presented down-core pre-industrial  $\delta^{13}\text{C}_{\text{DINO}}$  shows a similar mean, variance and data distribution to the nearby core-top samples (Supplement Fig. S1), suggesting that, at least for these nearby localities, the analyzed core-top specimens represent pre-industrial conditions. We find that an influence of the Suess effect and increased  $p\text{CO}_2$  impacts on the  $\delta^{13}\text{C}_{\text{DINO}}$  data is the most likely factor to explain the appearance of a small number of predominantly  $^{13}\text{C}$ -depleted outliers and resulting (subtle) negative skewing of the  $\delta^{13}\text{C}$  distributions (Fig. 4).

We use the background and outlier-corrected  $\delta^{13}\text{C}_{\text{DINO}}$  data and compare these data with  $\text{CO}_2$  conditions prevalent around 1850 CE (Fig. 6) and explore the effects of age inte-

gration by propagating different error distributions representative of the  $p\text{CO}_2$  change since 1850 CE (Fig. 7b–d). For practical purposes, we assume all  $\delta^{13}\text{C}_{\text{DINO}}$  populations to be normally distributed for further statistical analyses. We then use the mean carbon isotope value ( $\delta^{13}\text{C}_{\text{DINO}}$ ) and signal intensity in volt seconds (Vs) of each sample. The standard error of the mean ranges from  $\sim 0.2\%$  to  $0.7\%$  and depends primarily on the number of measurements in cases where  $n < 30$ , in line with expected values based on replicate measurements of the PE standard (van Roij et al., 2017) (Fig. 5).

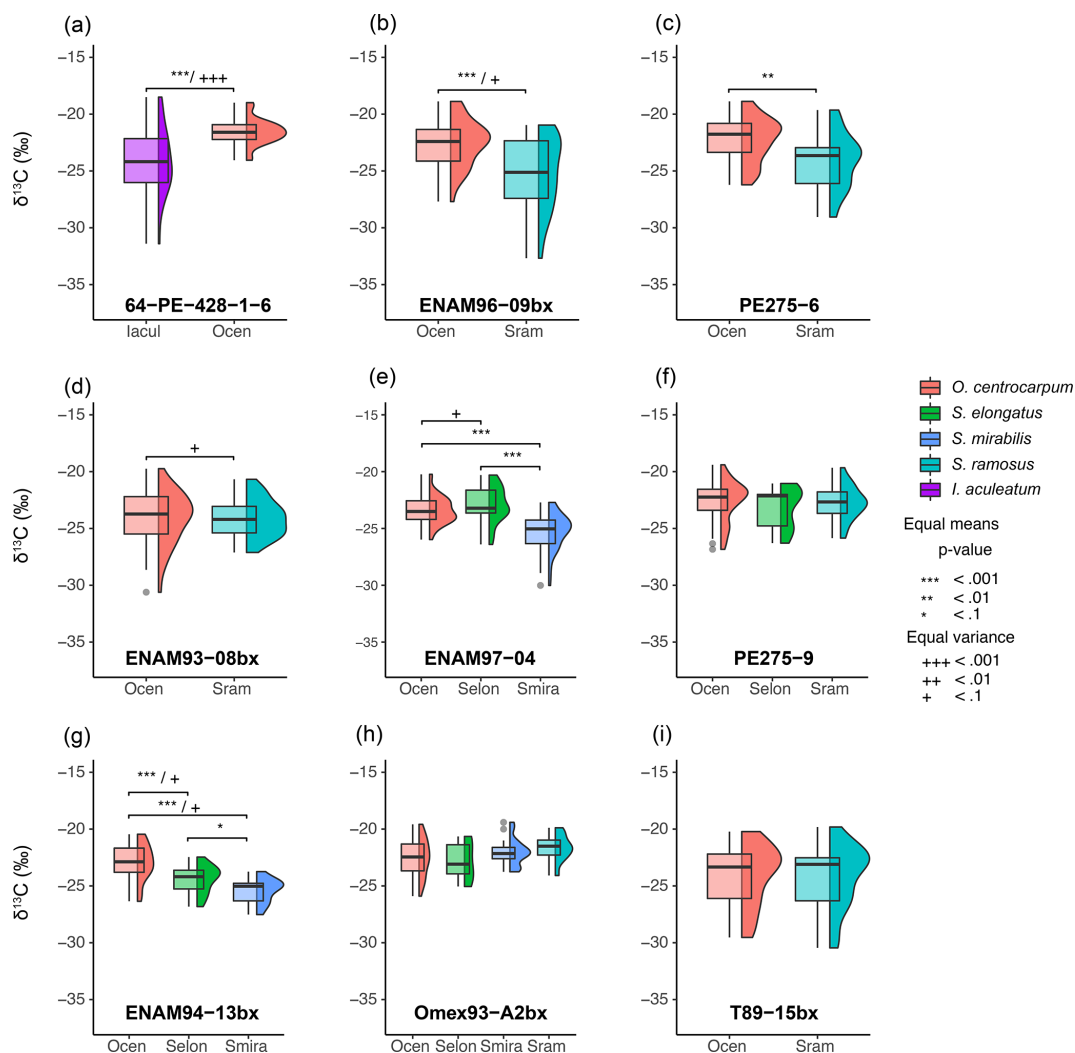
### 3.3 Environmental parameters and correlation

The range of measured  $\delta^{13}\text{C}_{\text{DINO}}$  values ( $\sim 5\%$ ) far exceeds the variability in surface ocean  $\delta^{13}\text{C}_{\text{DIC}}$  ( $\sim 1\%$ ) and  $\delta^{13}\text{C}_{\text{CO}_2}$  ( $\sim 2.5\%$ ), implying the observed range likely reflects differences in fractionation related to changing uptake or leakage of different inorganic carbon phases ( $\text{CO}_2$  and  $\text{HCO}_3^-$ ; Sharkey and Berry, 1985; Hoins et al., 2016a), and this hence determines most of the variability in the  $\delta^{13}\text{C}_{\text{DINO}}$  data. Here, we quantitatively assess fractionation as a function of several environmental parameters.

The simple (non-weighted) linear regressions show poor correlations between  $\epsilon_{p\text{-DIC}}$  and  $\delta^{13}\text{C}_{\text{DIC}}$  for all environmental parameters, and the correlations slightly improve when compared to  $\epsilon_{p\text{-CO}_2}$  (Tables 2, 3). However, none of the tested parameters individually explain the majority of the observed variance in  $\epsilon_{p\text{-DIC}}$  (maximum  $R^2$ ;  $\sim 0.1$ ) or  $\epsilon_{p\text{-CO}_2}$  (maximum  $R^2$  with  $p\text{CO}_2$ ;  $\sim 0.38$ ), despite high significance (low  $p$  values) of the regressions. The explained variance increases when polynomial regressions are applied. Several controlled growth experiments indeed show a non-linear response of  $\epsilon_p$  as a function of  $p\text{CO}_2$  of the growth medium (Hoins et al., 2015), although the number of data points in such experiments limits full mathematical descriptions of the trends within the  $p\text{CO}_2$  range of this field study. Here, a second-order polynomial (quadratic) regression achieves an  $R^2$  of  $\sim 0.74$  and  $\sim 0.79$  for the non-weighted and weighted versions, respectively.

It is conceivable that other environmental parameters also significantly contribute to  $\epsilon_p$  variability (Fig. 6). For example,  $[\text{PO}_4^{3-}]$ ,  $[\text{NO}_3]$  and  $p\text{CO}_2$  contribute significantly to a (linear) multiple-regression model, which takes the form of  $\epsilon_{p\text{-CO}_2} = c + x\text{CO}_2 + y\text{PO}_4 + z\text{NO}_3$ , where  $c$ ,  $x$ ,  $y$  and  $z$  are numerical constants. The multiple-regression model using these three parameters covers  $\sim 58\%$  of the variance in *O. centrocarpum*  $\epsilon_{p\text{-CO}_2}$  (not weighted) and  $67\%$  when weighted to number of measurements per sample. Including more parameters, such as SST, oxygen concentrations or other carbonate system parameters, explains slightly more of the observed variance but does not significantly improve the model. The residual mean standard error (RSME) of the carbon–nitrogen–phosphorus (CNP)- $\epsilon_p$  multiple-regression model is  $\sim 1.45\%$  while a linear regression with only  $p\text{CO}_2$





**Figure 4.** Carbon isotope measurements for multiple species; each panel represents a single sample after eliminating extremely small measurements sizes ( $< 0.2$  Vs), carrying out background correction and removing outliers ( $\pm 2.5$  interquartile range) (Sect. 3.2). Each box-whisker and  $\delta^{13}\text{C}$  distribution plot represents a set of measurements for a single species at their respective locality; note that tailing towards negative  $\delta^{13}\text{C}$  is common. Brackets above species  $\delta^{13}\text{C}$  populations indicate significant differences in means and variance for different species within a single sample.

yields 1.7‰. Only weighted regressions are given here and reported ranges of the constants represent 1 standard error or equivalent percentiles in the case of Monte Carlo-simulated errors. These models have the following optimal formats:

– Equation (1), linear.

$$\varepsilon_{p\text{-CO}_2} = 6.6 \pm 2.1 + 0.031 \pm 0.008 p\text{CO}_2 \quad (1)$$

(adjusted  $R^2 = 0.48$ ,  $p = 0.001$ ,  $\text{RSME} = 1.7$ ‰) (Fig. 6b).

– Equation (2a), quadratic (data without error propagation and only suitable for use  $> 240 \mu\text{atm}$ ).

$$\varepsilon_{p\text{-CO}_2} = 40.8 \pm 7.2 - 0.23 \pm 0.055 p\text{CO}_2 + 4.88 \times 10^{-4} \pm 1 \times 10^{-4} p\text{CO}_2^2 \quad (2a)$$

(adjusted  $R^2 = 0.79$ ,  $p < 0.001$ ,  $\text{RSME} = 1.13$ ‰) (Fig. 6b).

– Equation (2b), quadratic (Monte Carlo-constrained errors – analytical for  $p\text{CO}_2$  and  $\varepsilon_{p\text{-CO}_2}$ ) (Fig. 7b).

$$\varepsilon_{p\text{-CO}_2} = 35.6^{+5.8}_{-5.6} - 0.19^{+0.045}_{-0.045} p\text{CO}_2 + 4.1^{+0.91}_{-0.88} \times 10^{-4} p\text{CO}_2^2. \quad (2b)$$

– Equation (2c), quadratic (as Eq. 2b with additional  $45 \pm 15 \text{ ppm } p\text{CO}_2$  error) (Fig. 7c).

$$\varepsilon_{p\text{-CO}_2} = 39.3^{+11.5}_{-8.8} - 0.19^{+0.058}_{-0.076} p\text{CO}_2 + 3.4^{+1.3}_{-0.95} \times 10^{-4} p\text{CO}_2^2. \quad (2c)$$

**Table 2.** Linear regression coefficients and significance for all samples where *O. centrocarpum* was analyzed ( $n = 19$ ), with  $\varepsilon_{\text{p-CO}_2}$  as dependent variable.

	Coeff.	SE	$t$	$p$	$R^2$
CO <sub>2</sub> (mol kg <sup>-1</sup> )	$5.603 \times 10^5$	−0.447	0.66039	0.01163	
CO <sub>3</sub> <sup>2-</sup> (mol kg <sup>-1</sup> )	41 263.663	15 432.856	2.674	0.0160	0.296
HCO <sub>3</sub> <sup>-</sup> (mol kg <sup>-1</sup> )	809.88	7057.32	0.115	0.910	0.0007741
DIC (mol kg <sup>-1</sup> )	5655.859	6266.998	0.902	0.379	0.04572
SST (°C)	0.16971	0.06043	2.809	0.0121	0.3169
SSS (psu, practical salinity unit)	0.6737	0.3604	1.869	0.0789	0.1705
PO <sub>4</sub> <sup>3-</sup> (μmol kg <sup>-1</sup> )	−5.5131	2.1102	−2.613	0.0182	0.2865
NO <sub>3</sub> <sup>-</sup> (μmol kg <sup>-1</sup> )	−0.4484	0.2493	−1.798	0.0899	0.1599
Si (μmol kg <sup>-1</sup> )	−0.2553	0.3298	−0.774	0.449	0.03405
O <sub>2</sub> (mL L <sup>-1</sup> )	−1.274	0.433	−2.943	0.0091	0.3375
ALK (alkalinity, mol kg <sup>-1</sup> )	7621.253	4712.178	1.617	0.124	0.1334
$p\text{CO}_2 \sim 1850$	0.024945	0.007752	3.218	0.005050	0.3785

**Table 3.** As Table 2 but with  $\varepsilon_{\text{p-DIC}}$  as dependent variable.

	Coeff.	SE	$t$	$p$	$R^2$
CO <sub>2</sub> (mol kg <sup>-1</sup> )	$1.133 \times 10^5$	$4.760 \times 10^5$	0.238	0.815	0.00332
CO <sub>3</sub> <sup>2-</sup> (mol kg <sup>-1</sup> )	19 589.987	14 818.493	1.322	0.204	0.09322
HCO <sub>3</sub> <sup>-</sup> (mol kg <sup>-1</sup> )	6547.21	5758.02	1.137	0.271	0.07068
DIC (mol kg <sup>-1</sup> )	7803.835	5086.841	1.534	0.143	0.1216
SST (°C)	0.04959	0.06068	0.817	0.425	0.03781
SSS (psu)	0.4043	0.3201	1.263	0.224	0.08579
PO <sub>4</sub> <sup>3-</sup> (μmol kg <sup>-1</sup> )	−2.3993	2.0318	−1.181	0.254	0.07581
NO <sub>3</sub> <sup>-</sup> (μmol kg <sup>-1</sup> )	−0.05666	0.22969	−0.247	0.808	0.003567
Si (μmol kg <sup>-1</sup> )	−0.01925	0.28388	−0.068	0.947	0.0002703
O <sub>2</sub> (mL L <sup>-1</sup> )	−0.4170	0.4385	−0.951	0.355	0.05051
ALK (mol kg <sup>-1</sup> )	6754.843	3956.586	1.707	0.106	0.1464
$p\text{CO}_2 \sim 1850$	0.010083	0.007952	1.268	0.222	0.0864

– Equation (2d), quadratic (as 2b with resampled  $p\text{CO}_2$  rise 1800–2000 CE) (Fig. 7d).

$$\varepsilon_{\text{p-CO}_2} = 29.8^{+11.0}_{-8.0} - 0.13^{+0.061}_{-0.084} p\text{CO}_2 + 2.6^{+1.5}_{-1.1} \times 10^{-4} p\text{CO}_2^2. \quad (2d)$$

– Equation (3a), multiple-regression linear (Fig. 6e).

$$\varepsilon_{\text{p-CO}_2} = 6.0 \pm 3.1 + 0.034 \pm 0.01 p\text{CO}_2 + 1.22 \pm 0.47 \text{NO}_3 - 10.85 \pm 3.7 \text{PO}_4^{3-} \quad (3a)$$

(adjusted  $R^2 = 0.67$ ,  $p < 0.001$ , RSME = 1.45 ‰).

– Equation (3b) adjusted for application in the paleo-domain.

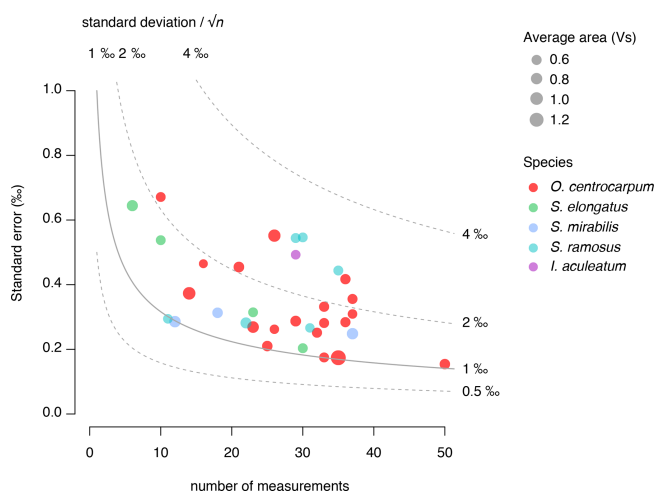
$$\varepsilon_{\text{p-CO}_2} = 6.0 \pm 3.1 + 0.034 \pm 0.01 p\text{CO}_2 - 1.1 \pm 5.3 \text{PO}_4^{3-}. \quad (3b)$$

– Equation (4), multiple-regression linear.

$$\varepsilon_{\text{p-DIC}} = 18.4 \pm 3.1 + 0.025 \pm 0.01 p\text{CO}_2 + 1.45 \pm 0.47 \text{NO}_3 - 11.1 \pm 3.7 \text{PO}_4^{3-} \quad (4)$$

(adjusted  $R^2 = 0.52$ ,  $p = 0.01$ , RSME = 1.44 ‰).

The two linear multiple-regression models are offset (Eqs. 3a and 4), primarily because of the carbon isotope fractionation between HCO<sub>3</sub><sup>-</sup> and CO<sub>2</sub>. The slope with respect to  $p\text{CO}_2$  also varies slightly between the models for  $\varepsilon_{\text{p-DIC}}$  and  $\varepsilon_{\text{p-CO}_2}$  due to the temperature-dependent fractionation between HCO<sub>3</sub><sup>-</sup> and CO<sub>2</sub>, but the slopes with NO<sub>3</sub><sup>-</sup> and PO<sub>4</sub><sup>3-</sup> are indistinguishable. The quadratic regression seemingly better fits the variability observed in  $\varepsilon_{\text{p-CO}_2}$  compared to other (multiple) linear regressions and removes any structure in the residuals, potentially signaling a non-linear response in  $\varepsilon_{\text{p-CO}_2}$  to  $p\text{CO}_2$ . The quadratic regression also indicates insensitivity to  $p\text{CO}_2 \leq 240 \mu\text{atm}$  and should not be used below this value (Fig. 6b). The Monte Carlo simulations



**Figure 5.** Relation of standard error of  $\delta^{13}\text{C}_{\text{DINO}}$  (‰) with the number of measurements and signal intensity (area in volt seconds (Vs)). Colors correspond to the various analyzed species.

of scenarios where an additional  $p\text{CO}_2$  uncertainty is imposed as a normally distributed mixture of pre-industrial and modern values, offsetting  $p\text{CO}_2$  by  $+45 \pm 15$  ppm (Fig. 7c, Eq. 2c), and a resampled uncertainty derived from the  $p\text{CO}_2$  rise since pre-industrial times (Fig. 7d, Eq. 2d) shows that the parameters of the quadratic regression are fairly robust to these uncertainties (i.e., none of the parameters become insignificant at  $p > 0.05$ ), although the absolute  $p\text{CO}_2$  values and errors increase.

## 4 Discussion

### 4.1 Absolute values, comparison to marine organic matter

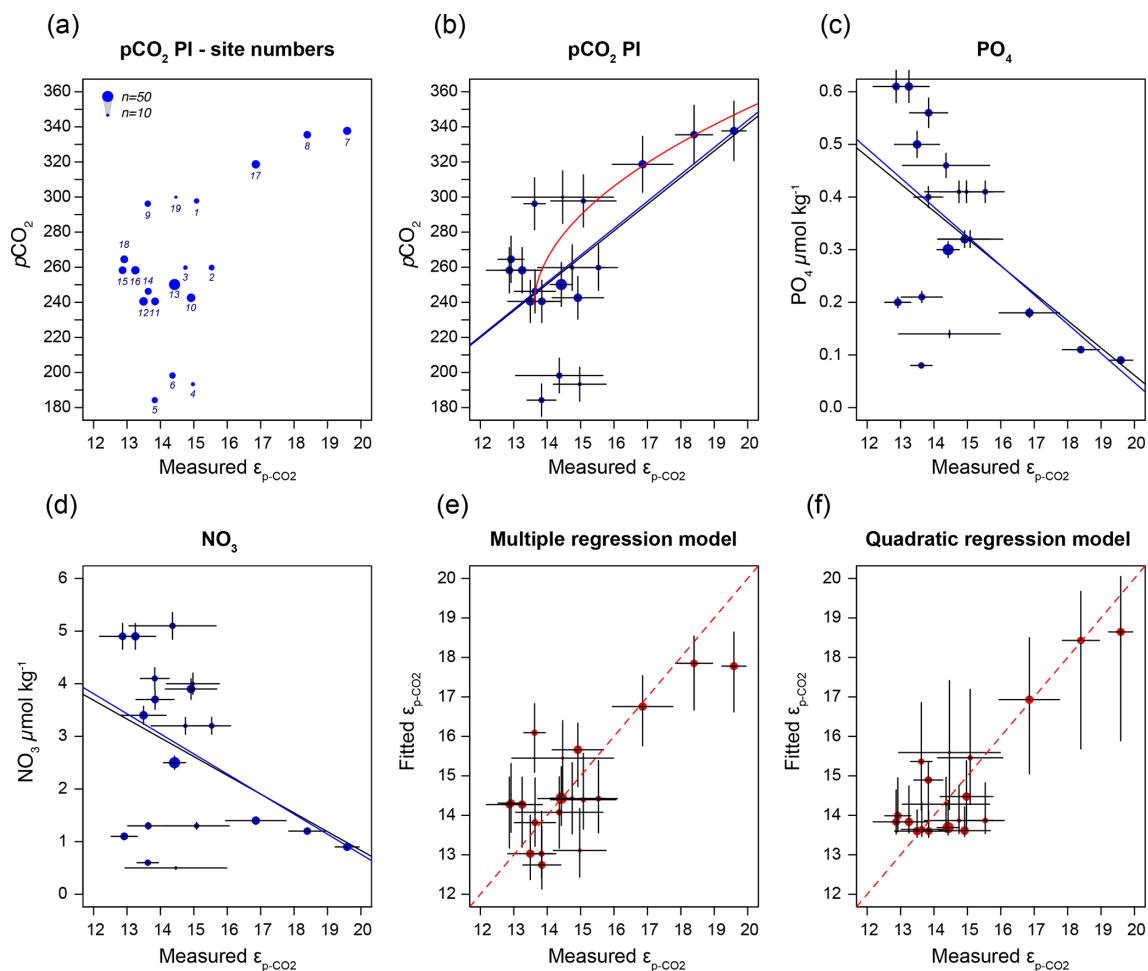
The recorded  $\delta^{13}\text{C}_{\text{DINO}}$  range and absolute values ( $\sim -18$ ‰ to  $-35$ ‰) correspond well with global  $\delta^{13}\text{C}$  values previously reported for marine particulate organic matter ( $\delta^{13}\text{C}_{\text{POC}}$ ) (e.g., Freeman and Hayes, 1992; Goericke and Fry, 1994) and modeled phytoplankton biomass (e.g., Magozzi et al., 2017; Tagliabue and Bopp, 2008). Consequently,  $\epsilon_{\text{p-DIC}}$  and  $\epsilon_{\text{p-CO}_2}$  are also within the expected range for general marine particulate organic matter. However, the intra-sample variance of  $\delta^{13}\text{C}_{\text{DINO}}$  appears to be substantial, often spanning most of the full range ( $\sim 10$ ‰) observed for  $\delta^{13}\text{C}_{\text{POC}}$ . Some of the observed variability might be related to the limited analytical precision during measurements of the extremely small numbers of carbon of individual dinocysts. Fully constraining the contribution of this analytical uncertainty to the observed variance is, however, not possible because of unresolvable micrometer-scale heterogeneity in the PE standard (van Roij et al., 2017; Sluijs et al., 2018). In most cases, the variance in  $\delta^{13}\text{C}_{\text{DINO}}$  is similar to that of the standard. Still, it is likely that some of

the seasonal  $\delta^{13}\text{C}_{\text{DIC}}$  differences are also recorded in the  $\delta^{13}\text{C}_{\text{DINO}}$  and that some additional inter-specimen  $\delta^{13}\text{C}$  variance is present. This is to be expected since the  $\delta^{13}\text{C}$  populations from our integrated core-top samples span seasons and decades and thus also considerable variability in seawater properties and population change. In addition, growth-induced randomness and changes in  $\delta^{13}\text{C}$  and DIC in the cell's microenvironment likely contributed to inter-specimen variability. Note that in our data inter-specimen variability is still underestimated because we analyzed three to seven specimens per ablation event, as single-cyst carbon yield ( $\sim 7$  ng C) from these cyst-sizes approached the limit for reliable measurements (van Roij et al., 2016). We minimized potential influence of differences in cell size or shape through manual selection. We thus analyzed a population where the pre-selection of similar-sized cysts restricts the variance in cell surface area and volume, unlike biomarker-based proxies for which the cell size has to be reconstructed independently (Henderiks and Pagani, 2007; Stoll et al., 2019). This approach could reduce scatter in the relation of  $\epsilon_{\text{p}}$  to environmental variables (Popp et al., 1998).

### 4.2 Cell–cyst offset

One of the striking differences between the here generated data and the existing culture experiments is that carbon isotope fractionation of dinocysts in the natural environment appears to be much larger than that of motile cells from controlled growth (dilute batch) experiments (Hoins et al., 2015, 2016b). We find average  $\epsilon_{\text{p}}$  values ranging between  $13$ ‰– $20$ ‰ and  $23$ ‰– $29$ ‰ with respect to  $\text{CO}_2$  and DIC. Cultured cells of *O. centrocarpum* yielded not only a smaller overall  $\epsilon_{\text{p}}$  but also a smaller range ( $\sim 8$ ‰– $12$ ‰ and  $18.5$ ‰– $20$ ‰) across a larger  $\text{CO}_2$  gradient, implying the cysts have a much steeper fractionation slope with  $\text{CO}_2$  compared to the motile cells. Despite these differences, the average  $\epsilon_{\text{p}}$  for *Spiniferites* species (*S. ramosus*, *elongatus* and *mirabilis*) is often somewhat larger than for *O. centrocarpum* (Fig. 4). This is consistent with culture experiments that showed larger  $\text{CO}_2$  dependency and overall slightly larger  $\epsilon_{\text{p}}$  in the motile species *G. spinifera* compared to *P. reticulatum* (Hoins et al., 2015).

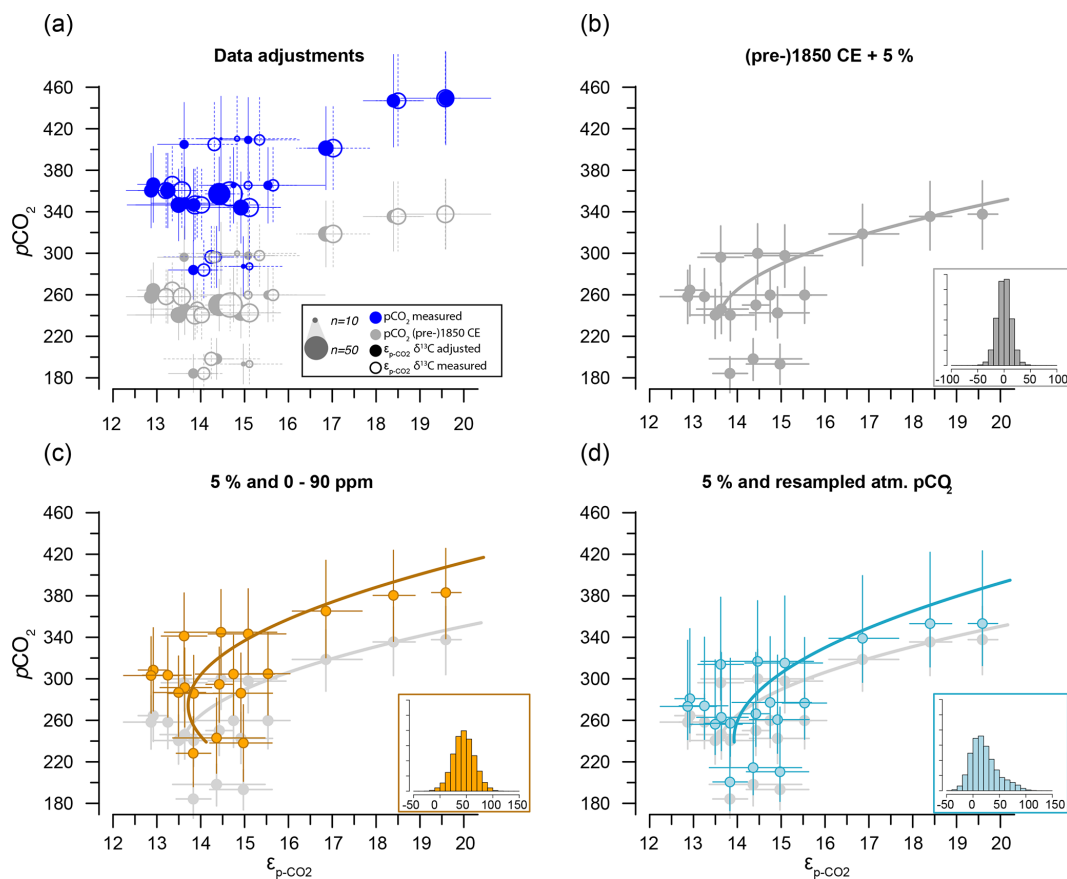
While the cultured single strains and dinoflagellate populations in nature may behave somewhat differently, we do not expect that this alone underlies such a marked offset between the motile cultured cells and natural cysts. Natural cysts and cultured cells seem consistently offset in  $\delta^{13}\text{C}$ , although at present the exact amplitude of this offset cannot be determined. However, such an offset is in line with certain compounds in dinoflagellate cells being depleted in  $^{13}\text{C}$  relative to the bulk biomass (Schouten et al., 1998; Wilkes et al., 2017). The organic-walled dinocysts consist of mostly aliphatic and aromatic compounds, forming a complex biopolymer referred to as dinosporin (de Leeuw et al., 2006; Versteegh et al., 2007, 2012). Depending on the



**Figure 6.** (a) Sample-specific  $\varepsilon_{p\text{-CO}_2}$  of *O. centrocarpum* relative to pre-industrial  $p\text{CO}_2$ ; numbers correspond to localities listed in Table 1. (b) Regression analyses for  $\varepsilon_{p\text{-CO}_2}$  of *O. centrocarpum* relative to  $p\text{CO}_2$  (measured and corrected to pre-industrial values); the black line represents simple linear regression, the blue line represents weighted linear regressions and the red line represents weighted quadratic regression. (c) Phosphate concentrations ( $\text{PO}_4^{3-}$ ). (d) Nitrate concentrations ( $\text{NO}_3^-$ ). (e) Fitted values illustrating the multiple-regression model performance using parameters (a)–(c) relative to measured  $\varepsilon_{p\text{-CO}_2}$ . (f) Fitted values using only pre-industrial  $p\text{CO}_2$  but applying a quadratic regression (red curve in panel b). Errors in panels (b)–(d) represent 5% of the measured value and errors on the fitted values in panels (d) and (e) represent propagated errors in both measurements and environmental variables (as shown in panel b–d) using Monte Carlo simulations ( $n = 1000$ ) for regression models. Symbol size (top left corner of panel a) represents the number of measurements within each sample.

biosynthetic pathway of the cyst material and the derivation or degradation of the original compounds, this may result in offsets in  $\delta^{13}\text{C}$  values between cysts and the motile cells. A potential additional fractionation might be introduced during taphonomy and also later by the processing of sediments to concentrate the dinocysts from sediment samples. The sediment processing involves hydrochloric and hydrofluoric acids, which affects the more labile organic compounds. Last, it is conceivable that fractionation in the dilute batch experiments may be reduced by, e.g., higher-than-natural growth rates. This may be supported by chemostat culture experiments on *Alexandrium tamarense* (Wilkes et al., 2017), which show a (much) greater fractionation compared to the

dilute batch experiments (Hoins et al., 2015). However, the enhanced fractionation recorded in chemostat experiments is likely an artifact of isotope equilibration times exceeding  $\text{CO}_2$  uptake rates (Brandenburg et al., 2022; Zhang et al., 2022). The range of options cannot be narrowed down until cultured cysts are compared to their motile cells harvested from the same culture and treated with similar techniques as used for the sediments. Until these data become available, inferences on the origin and amplitude of the offsets between the cells and cysts of *O. centrocarpum* and *Spiniferites* remain speculative.



**Figure 7.** Sensitivity tests on the potential effect of anthropogenic carbon emissions. **(a)** Effects of data treatment on the difference between measured and adjusted  $p\text{CO}_2$  and  $\varepsilon_{p\text{-CO}_2}$  (same as Fig. 3). Open symbols indicate measured  $\delta^{13}\text{C}$ ; closed symbols represent data after eliminating small signals ( $< 0.2$  Vs) and outliers. Blue dots represent measured  $\text{CO}_2$  values and gray dots indicate the  $\text{CO}_2$  around 1850 CE. **(b)** Quadratic regression (red line in Fig. 6b) with propagated analytical error in  $p\text{CO}_2$  and  $\delta^{13}\text{C}$  only, using  $\text{CO}_2$  values around 1850 CE (gray filled symbols in panel a). **(c)** As in **(b)** but with the addition of a  $+45 \pm 15$  ppm error to reflect potential impact of anthropogenic  $\text{CO}_2$  in orange. Gray dots and curve of panel **(b)** are added as a comparison. **(d)** As in **(b)** but with the addition of a  $\text{CO}_2$  increase relative to pre-industrial values, sampled for the period 1800–2000 CE. Insets on the bottom right in panels **(b)**, **(c)** and **(d)** show the combined error distributions imposed on  $p\text{CO}_2$ . All error bars in panels **(b)**–**(d)** on  $p\text{CO}_2$  and  $\varepsilon_{p\text{-CO}_2}$  are 2.5%–97.5% percentile ranges from Monte Carlo simulations ( $n = 1000$ ).

### 4.3 Environmental controls on carbon isotope fractionation

Carbon isotope fractionation is determined by RuBisCO and several environmental parameters, dominantly  $p\text{CO}_2$  but also cell size and shape, growth rates, and nutrient or light regimes (e.g., Freeman and Hayes, 1992; Pagani, 2013; Popp et al., 1998; Stoll et al., 2019, and many others). In most cases, fractionation is  $p\text{CO}_2$ -dependent, and consequently  $\varepsilon_p$  in various groups and genera has been used as a paleo- $p\text{CO}_2$  proxy (e.g., Freeman and Hayes, 1992; Pagani, 2013; Rae et al., 2021). We performed a broad-spectrum multiple-regression analysis to identify environmental factors that contribute significantly to dinocyst  $\varepsilon_p$ . The most important parameter is  $p\text{CO}_2$ , in line with previous studies (Freeman and Hayes, 1992). A large part of the remaining variability

can be attributed to growth rate and ultimately nutrient content, specifically nitrate and phosphate of the surface waters, in line with previous studies on phytoplankton and dinoflagellate carbon isotope fractionation (Popp et al., 1998; Hoins et al., 2016b; Wilkes et al., 2017; Wilkes and Pearson, 2019). We find that  $p\text{CO}_2$  and  $\varepsilon_p$  are positively correlated, while  $\text{NO}_3^-$  and  $\text{PO}_4^{3-}$  are negatively correlated with  $\varepsilon_p$  (Fig. 6c, d).

The inclusion of nutrient levels as environmental factors in the carbon isotope fractionation reduces the error in the fit between the measured and modeled fractionation. In the rare cases where paleo-nutrient concentrations are estimated, reconstructions usually cover only  $[\text{PO}_4^{3-}]$ , implying the  $[\text{NO}_3^-]$  in Eqs. (3a) and (4) has to be approximated from  $[\text{PO}_4^{3-}]$ . As  $[\text{PO}_4^{3-}]$  and  $[\text{NO}_3^-]$  are generally well-correlated (typically  $\sim 1 : 10$   $[\text{PO}_4^{3-}] : [\text{NO}_3^-]$ ; here  $\sim 1 : 8$ ) in ocean wa-

ters, this is relatively straightforward. For application in the paleo-domain, however, nutrient ( $[\text{PO}_4^{3-}]$ ) reconstructions may not be available or may be considered too unreliable to provide meaningful constraints. In specific cases where accurate paleo- $[\text{PO}_4]$  estimates or large changes are reconstructed, Eq. (3b) may be applied, in which  $[\text{NO}_3]$  is substituted for  $[\text{PO}_4^{3-}]$  at a 1 : 8 ratio. However, unless there are clear changes in the nutrient regime or sufficient proxy constraints on the nutrient concentrations, we prefer a calibration based exclusively on carbon isotope fractionation and carbonate system parameters, including sea surface temperature to calculate  $\delta^{13}\text{C}_{\text{CO}_2}$  from  $\delta^{13}\text{C}_{\text{DIC}}$  to reconstruct  $p\text{CO}_2$ .

#### 4.4 Influence of carbon-concentrating mechanisms: $\text{CO}_2$ insensitivity

As many phototrophs, including dinoflagellates, have mechanisms for concentrating  $\text{CO}_2$  near the cell membrane, the sensitivity of carbon isotope fractionation to ambient  $\text{CO}_2$  is expected to diminish below a certain concentration (M. R. Badger, 2003; M. P. S. Badger, 2021; Stoll et al., 2019). This is particularly the case as dinoflagellates utilize type II RuBisCO, which has generally poorer performance compared to type I RuBisCO at low  $\text{CO}_2$  concentrations (Giordano et al., 2005). Indeed, substantial activity of the carboxyl anhydrase (CA) enzyme, which facilitates the conversion of  $\text{HCO}_3^-$  to  $\text{CO}_2$  inside the cell or near the membrane, was observed in numerous dinoflagellate species, including *Lingulodinium* (Lapointe et al., 2008), *Symbiodinium* (Leggat et al., 1999) and the here analyzed *Operculodinium* (Ratti et al., 2007). Also, here we find a relatively low sensitivity at the lower end of the  $\text{CO}_2$  scale. The lower  $\text{CO}_2$  values correspond to the northern-most locations, with trends below  $240\ \mu\text{atm}$  becoming somewhat obscured, at a minimum  $\varepsilon_{\text{p-CO}_2}$  around 13‰ (Fig. 6b). However, part of this leveling of the proxy relationship may reflect the locally higher nutrient concentrations offsetting the higher  $\text{CO}_2$ . Though growth rates have a clear influence on  $\varepsilon_{\text{p}}$  in algal groups (Burkhardt et al., 1999), including dinoflagellates (Wilkes et al., 2017), the dilute batch culturing experiments conducted with *P. reticulatum* showed no clear influence of growth rates on  $\varepsilon_{\text{p}}$  (see also Sect. 4.2). It is also conceivable that higher growth rates influence  $\varepsilon_{\text{p}}$  indirectly through, for example, seasonally enhanced  $\text{CO}_2$  drawdown, resulting in higher  $\delta^{13}\text{C}$  values in the remaining DIC. This effect may be enhanced by the relatively short growing season at the high latitudes. However, in culture experiments at low  $\text{CO}_2$  concentrations with other dinoflagellate species,  $^{13}\text{C}$  fractionation was higher under nutrient-limiting conditions than under replete conditions (Hoins et al., 2016b). Because of these confounding factors, the influence of carbon-concentrating mechanisms on  $\varepsilon_{\text{p}}$  in *O. centrocarpum* is difficult to gauge with the presently available data and would ideally be tested using high nutrient or very low  $\text{CO}_2$  concentrations.

Still, also in the relatively limited range the current ocean offers for testing  $p\text{CO}_2$  proxies we have established a robust, albeit not overly sensitive, relation between  $p\text{CO}_2$  and dinocyst  $\delta^{13}\text{C}$ . Our cyst-based calibration yields more conservative and arguably more realistic absolute  $\text{CO}_2$  estimates and variability compared to available culture-based calibrations as it is based on the same compounds as will be analyzed in the paleo-domain. However, the low sensitivity at low  $\text{CO}_2$  implies that, until better constraints become available, the proposed calibration is potentially less suitable for application across, for example, the Pleistocene glacial periods. Further, it is important to realize that the value of  $240\ \mu\text{atm}$  is based on the assumption that the  $\varepsilon_{\text{p-CO}_2}$ - $p\text{CO}_2$  relation originated from cysts that have not been affected by the Suess effect and thus represent a lower limit for  $\text{CO}_2$  (in)sensitivity. While our data do not preclude fractionation smaller than the here observed minimum ( $\sim 13\text{‰}$ ) during low- $p\text{CO}_2$  periods, increased sensitivity at higher  $\text{CO}_2$  suggests that  $\text{CO}_2$  above (minimum)  $240\ \mu\text{atm}$  and  $\text{CO}_2$  variability can be reconstructed with reasonable precision.

#### 4.5 Challenges of age control and potential caveats associated with anthropogenic carbon

A topic that warrants specific attention is the potential impact of anthropogenic carbon emissions on shaping the relation between  $\delta^{13}\text{C}_{\text{DINO}}$  and  $\varepsilon_{\text{p}}$ . Here, we assume that our outlier analyses preferentially excluded samples that were significantly affected by anthropogenic  $\text{CO}_2$  and that, on average, the remainder of our core-top  $\delta^{13}\text{C}_{\text{DINO}}$  was not appreciably affected by either decreasing  $\delta^{13}\text{C}_{\text{DIC}}$  or elevated  $p\text{CO}_2$ . The similarity in pre-industrial down-core  $\delta^{13}\text{C}_{\text{DINO}}$  and that of three core-top localities in the North Atlantic corroborate the validity of this assumption locally, but these observations cannot be extrapolated to other regions. Unfortunately, sedimentation rates or other constraints for cyst production datums are not available. Further, if sedimentation rates were available for core-top localities, that would constitute an imperfect solution as it cannot provide the required dinocyst-specific age distribution needed to obtain an appropriate local  $p\text{CO}_2$  correction to the datum of cyst production. This challenge may be unique to data such as those presented here, as studies on other substrates with  $p\text{CO}_2$ -proxy potential either could not generate data for individual single-celled organisms or have avoided the issue through other means, such as culture experiments (e.g., Pagani et al., 2002; Henehan et al., 2013), approaches that have other drawbacks.

Although we have no constraints on the ages of the cysts analyzed here, we can provide a meaningful test of the potential uncertainty added by our assumption that cysts are representative of pre-industrial conditions (Eq. 2a, b). The two scenarios that we have explored through Monte Carlo simulations show that, depending on the error distribution imposed on the assumed  $p\text{CO}_2$  used in the quadratic regressions (Eq. 2a, b), the fitted regression shifts towards higher

values (Fig. 7c) and may be steeper (Fig. 7d). These error-propagation case studies illustrate that our proposed pre-industrial transfer function (Eq. 2a, b), if it does indeed contain a substantial proportion of very recent dinocysts, is likely to lead to underestimated  $p\text{CO}_2$  and perhaps  $p\text{CO}_2$  variability when applied in the paleo-domain. Consequently, we recommend future studies target, for example, sediment trap collections and culture-derived dinocysts to validate the results obtained here.

## 5 Proxy potential, limitations and calibrations

Our new modern ocean single-species carbon isotope fractionation dataset shows promising trends with environmental variables,  $p\text{CO}_2$  and nutrients. The selection of individual cysts allows control of cell size and species, which reduces uncertainty in proxy calibration and application compared to approaches based on organic substrates, which inevitably integrate entire communities. Although this approach has clear benefits, it also poses a unique challenge as the impact of anthropogenic carbon emissions on individual single-celled organisms must be considered. Based on our analyses, we expect this to be a relatively minor factor. In a worst-case scenario, however, we find that, although a helpful simplification, the assumption that all dinocysts from the core-top samples represent pre-industrial conditions may lead to an underestimate of  $p\text{CO}_2$  and perhaps also  $p\text{CO}_2$  variability when applied as a proxy in the paleo-domain.

In addition, many of the challenges associated with other proxies based on organic substrates are encountered here as well. For example, like in cultures (Hoins et al., 2016b), we observed an impact of nutrients on carbon isotope fractionation, possibly related to differences in growth rates. Similarly, at low  $p\text{CO}_2$  values sensitivity is reduced, possibly because of carbon-concentrating mechanisms involved in dinoflagellate C uptake, as observed in culturing experiments (e.g., Hoins et al., 2016a). Another remaining challenge is the observed difference between the cultured populations and cysts from the core-top sediments. This is a pronounced difference, not only in the absolute isotope fractionation values but also in the slope of the  $\text{CO}_2$  sensitivity, which appears to be much larger for the cysts and requires attention in future culture studies.

The offset in  $\delta^{13}\text{C}$ , combined with uncertainties in fractionation between the motile cells and dinocysts, implies that  $\text{CO}_2$  reconstructions using culture-based calibrations are more likely to overestimate past  $p\text{CO}_2$ . Furthermore, the large spread in our data ( $\sim 5\%$  between high and low  $\text{CO}_2$ ) suggests that, due to this high sensitivity in the cysts, the method is also suited to study population dynamics.

*Data availability.* All newly generated data will be available via a permanent online repository (Mendeley data <https://doi.org/10.17632/z6285myxkm.2>, Frieling, 2023) upon publication.

*Supplement.* The supplement related to this article is available online at: <https://doi.org/10.5194/bg-20-4651-2023-supplement>.

*Author contributions.* AS and GJR designed the study. LvR, IK and JF processed samples, generated and analyzed data. JF wrote the original draft. AS and GJR reviewed and edited the paper. AS acquired funding for this study.

*Competing interests.* The contact author has declared that none of the authors has any competing interests.

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