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PRICE INTEGRATION ALONG THE SPANISH GILTHEAD SEABREAM (SPARUS AURATA) VALUE CHAIN

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Global production of gilthead seabream (Sparus aurata) increased during the 2008 – 2013 period from 129 thousand tonnes valued 698 million USD in 2008 to 173 thousand tonnes valued 1065 million USD in 2013.



Global production of gilthead seabream



- Greece and Turkey are the world gilthead seabream leading producers with 42% and 21% of the volume and 45% and 17% of the value produced, respectively.
- The EU produced nearly 109 thousand tonnes, valued 706 million USD, in 2012, accounting for 63% of global production volume and 66% of the value.



Gilthead seabream leading producers

In the EU, the main European producer is Greece with 73.3 thousand tonnes, followed by Spain and Italy with around 18.9 and 5.4 thousand tonnes, respectively (FAO, 2016).







DS-045409-EU Trade Since 1988 by HS2, 4, 6 and CN8. Source EUROSTAT. 03026995 – Fresh or chilled gilt-head seabream "Sparus aurata"





Evolution of seabream imports prices in Spain



Source EUROSTAT



- **Sea bream aquaculture in the Mediterranean Sea faces various constraints that include:**
 - Increasing supply in fully developed markets
 - Effects of financial crises in many consumer countries
 - Difficulties of negotiating with **concentrated retailers**
- These constraints have raised interest in analyzing the ways in which prices are set in the international market and along the value chain
- Accordingly, the aims of this work are:
 - Analyse international competition in the Spanish seabream market depending on the country of origin. (Horizontal price integration analyses)
 - Analyse the price transmission mechanisms on the Spanish seabream market in order to know how the negotiation power is distributed along the value chain. (Vertical price integration analyses)



- Let is possible to study **price integration** by analyzing **price linkages**
- The price integration analysis has been used in various research applications in the field of fisheries markets:
 - Different levels of the value chain → Price transmission → Vertical price integration (Jimenez-Toribio et al. (2003); Guillotreau (2004); Jaffry (2005); Guillotreau et al. (2005); Asche et al. (2007) Guillen and Franquesa (2008); Jimenez-Toribio et al. (2010b); Sakai et al. (2012); Asche et al (2014).
 - Different products/producers → Market integration → Horizontal price integration (Nielsen (2004); Asche et al. (2005); Nielsen et al. (2007); Asche et al. (2007); Vinuya (2007); Nielsen et al. (2009); Jimenez-Toribio et al. (2010a); Asche et al. (2012); Rodriguez et al. (2013) Schrobback et al. (2014)).



- The limited sample size of the data used (2009 to 2013) did not allow conducting a single model. Instead, two separate models were used to illustrate horizontal and vertical price integration. Seasonality can not be assessed.
- The first model studied price competition among Greece, Turkey and Spain, the three largest producers of sea bream.
 - **Horizontal price integration** allows identification of who exerts price leadership in markets.
 - Ex-farm prices
 - Import prices
- A second model studied price transmission across ex-farm, wholesale and retail prices
 - Vertical price integration informs about the transmission of prices across value chain levels and potential issues of market power.
 - Ex-farm prices
 - Wholesale prices
 - Retail prices
- A new version of the second model studied transmission across imports from Turkey, wholesale and retail prices
 - Vertical price integration informs about the transmission of prices across value chain levels and potential issues of market power.
 - Turkey's import price
 - Wholesale prices
 - Retail prices



- Given the non-stationary nature of most of the price series, the statistical method used to study the relationships among these is the cointegration analysis.
- In the case of price analyses, these models are based on statistical techniques to capture price linkages, where some form of cointegration between prices defines the market and allows predicting prices and random changes in the price chain.
- Co-integration analysis requires non-stationary price series data and univariate unit root (nonstationary) test can indicate the stationary properties of the data (Norma Lopez et al. 2014)
- The Augmented Dickey-Fuller (ADF) test (Dickey & Fuller, 1979; 1981) is used to test the time series properties of the data (non-stationarity).

Model 1. linear trend	Z	Р	Model 2. no constant	Z	Р	Model 2.1. linear trend	z	Р
Greece (imports)	-1.996	0.602	Exfarm	0.111	0.717	Turkey (imports)	-1.675	0.762
Turkey (imports)	-1.675	0.762	Wholesaler	0.402	0.799	Wholesaler	-1.137	0.852
Spain (exfarm)	-1.897	0.655	Retailer	0.508	0.825	Retailer	-2.021	0.589

 All data series are Unit root I(1) at their levels. Unit root is rejected for the first differences in all price series.



- When the price series are non-stationary the Johansen test (Johansen, 1991) is the natural approach (Asche et al. 2007)
- The Johansen test corrects the problem of non-stationarity in the data series.
- Also weak exogeneity and Granger causality tests are applied to understand price leadership and price transmission

DATA

- The prices for sea bream at ex-farm, wholesale and retail levels have been collected weekly for sea bream from 2009 to 2013 by Spain's Ministry of Agriculture and Food through the Observatory of Food.
- Prices for Spain's imports from Greece and Turkey were obtained from the European Commission's Eurostat trade database.



MODEL 1. MARKET INTEGRATION ANALYSIS

The model involving the prices of imports from Greece and Turkey and the Spanish ex-farm prices resulted in **one cointegrating vector**.

Rank	Eigenvalue	Trace	р	Lmax	р
0	0.50548	59.503	0.0003	33.800	0.0022
1	0.35154	25.703	0.0504	20.791	0.0281
2	0.097261	4.9115	0.6155	4.9115	0.6171

This implicated that the three origins compete in the same delimitated market (Spain).

■Subsequent tests (Weak exogeneity test) indicated that Greek prices were endogenous, meaning that price changes for Greek are following the prices of Spanish and Turkey producers, which are exogenous.

Constrain	Statistic	р
a[1] = 0 (Greece)	7.54009	0.0060 (endogenous)
a[2] = 0 (Turkey)	0.323521	0.5695 (exogenous)
a[3] = 0 (Spain)	0.0320448	0.8579 (exogenous)



MODEL 1. MARKET INTEGRATION ANALYSIS

□ In order to obtain a more detailed information about the competitive relations between the origins in the market, we applied the **Granger causality test**.

Greece			
	All the lags from Greece	F(1. 54) = 1.1289 [0.2927]	
	All the lags from Turkey	F(1. 54) = 1.6458 [0.2050]	
	All the lags from Spain	F(1. 54) = 2.8205 [0.0988]	
Turkov			
типкеу			
	All the lags from Greece	F(1. 54) = 0.0259 [0.8727]	
	All the lags from Turkey	F(1. 54) = 5.8646 [0.0188]	
	All the lags from Spain	F(1. 54) = 1.6551 [0.2037]	
Spain			
	All the lags from Greece	F(1. 54) = 0.0372 [0.8478]	
	All the lags from Turkey	F(1. 54) = 2.2915 [0.1359]	+
	All the lags from Spain	F(1. 54) = 6.4896 [0.0137]	

Spanish and Turkish prices were found to be exogenous, caused by variables not included in the model, such as long-term contracts with local retailers or other institutions.



MODEL 2. PRICE TRANSMISSION ANALYSIS

The model assessing price transmission along the Spanish value chain also found one cointegrating vector.

Rank	Eigenvalue	Trace	р	Lmax	р
0	0.78397	101.36	0.0000	73.553	0.0000
1	0.43969	27.809	0.0000	27.805	0.0000
2	7.4558e-005	0.0035789	0.9758	0.0035789	0.9726

In this case, it confirmed that shifts in ex-farm price reached the retail level, suggesting that no actors exerted definitive market power.

□However, **ex-farm and retail prices were found to be endogenous**, and can be explained by the links among them and the wholesale level.

Constrain	Statistic	р
a[1] = 0 (Exfarm)	29.9492	4.43518e-008 (endogenous)
a[2] = 0 (Wholesaler)	0.0371	0.8473 (exogenous)
a[3] = 0 (Retailer)	14.7113	0.0001 (endogenous)

The **prices at the wholesale level**, instead, **were exogenous** and affected by causes not included in this model.



MODEL 2. PRICE TRANSMISSION ANALYSIS

□ In order to obtain a more detailed information about the price behavior along the value chain, we applied the **Granger causality test**.

Exfarm		
	All the lags from Exfarm	F(12. 12) = 9.6631 [0.0002]
	All the lags from Wholesaler	F(12. 12) = 6.4350 [0.0015]
	All the lags from Retailer	F(12. 12) = 9.2387 [0.0003]
Wholesaler		
	All the lags from Exfarm	F(12. 12) = 1.5538 [0.2282]
	All the lags from Wholesaler	F(12. 12) = 2.5647 [0.0582]
	All the lags from Retailer	F(12. 12) = 1.3029 [0.3270]
Retailer		
	All the lags from Exfarm	F(12. 12) = 2.8418 [0.0414]
	All the lags from Wholesaler	F(12. 12) = 2.8921 [0.0390]
	All the lags from Retailer	F(12. 12) = 171.93 [0.0000]

Domestic ex-farm prices were related to retail prices, but they did not seem to affect wholesale prices.



MODEL 2.1. PRICE TRANSMISSION ANALYSIS

Second estimation of the second model. To clarify the relationships among the different price series.

■When the prices of domestic ex-farm prices were substituted for Turkish imports , one cointegrating vector was found.

Rank	Eigenvalue	Trace	р	Lmax	р
0	0.64068	96.167	0.0000	49.130	0.0000
1	0.55673	47.037	0.0000	39.052	0.0000
2	0.15325	7.9850	0.2606	7.9850	0.2608

Once again, wholesale prices were exogenous, affecting both Turkish and retail prices.

This meant that **Turkish exporters were adopting the prices of Spanish wholesalers**, and **Spanish retailers adapted their prices as the wholesale price changed**.

Constrain	Statistic	р
a[1] = 0 (Turkey)	3.5532	0.0594 (endogenous at 90% sig. And exogenous at 95% sig.)
a[2] = 0 (Wholesaler)	0.8215	0.3647 (exogenous)
a[3] = 0 (Retailer)	3.6913	0.0547 (endogenous at 90% sig. And exogenous at 95% sig.)



MODEL 2.1. PRICE TRANSMISSION ANALYSIS

□ In order to obtain a more detailed information about the price behavior along the value chain, we applied the **Granger causality test**.

Turkey		
	All the lags from Turkey	F(12. 10) = 1.2384 [0.3725]
	All the lags from Wholesaler	F(12. 10) = 2.8789 [0.0518]
	All the lags from Retailer	F(12. 10) = 1.8319 [0.1728]
Wholesaler		
	All the lags from Turkey	F(12. 10) = 2.0395 [0.1338]
	All the lags from Wholesaler	F(12. 10) = 4.6161 [0.0107]
	All the lags from Retailer	F(12. 10) = 1.6639 [0.2139]
Retailer		
	All the lags from Turkey	F(12. 10) = 1.2993 [0.3439]
	All the lags from Wholesaler	F(12. 10) = 1.3075 [0.3402]]
	All the lags from Retailer	F(12. 10) = 2.3000 [0.0982]

Results suggest that retailer prices are exogenous and Turkish are endogenous (depend at

90% sig from the wholesalers)



Aim 1. Market delimitation

- As was expected, the Spanish seabream market is delimitated, and competitive.
- The ex-farm prices of sea bream in Spain are partially affected by international competition led by Turkish imports and by the prices set by wholesalers.
- Aim 2. Price transmission
 - Spanish farmers' prices interact with the retail level in a bidirectional relation indicating perfect price transmission between agents.
 - Retailers accommodate the prices of sea bream according to changes in farmers' costs, and also respond to changes in wholesale prices.
 - While the results in assessing the exercise of market power at any level of the value chain are not conclusive, there is evidence that Spanish wholesalers could exercise some degree of negotiation power.
 - While totally independent of the behavior of the other actors, wholesale prices for sea bream affected all other observed prices. All levels of the value chain adjusted their prices to changes in the wholesale level, whether for domestic or imported sea bream.



• We need more data.

We will be able to developed an extended model, which includes Turkey, Greece, Exfarm and Retailer.

This extended model will allow us:

1) Firstly, confirm or reject the results showed in the present work.

2) We will be able to apply more complex models







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THANK YOU

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