

### Effects of exclusive breastfeeding promotion interventions on child outcomes: a systematic review and meta-analysis

DIB, Sarah, FAIR, Frankie <a href="http://orcid.org/0000-0001-7613-3393">http://orcid.org/0000-0001-7613-3393</a>, MCCANN, Lucy Jane, NICHOLLS, Antonia, KALEA, Anastasia Z, SOLTANI, Hora <a href="http://orcid.org/0000-0001-9611-6777">http://orcid.org/0000-0001-9611-6777</a> and FEWTRELL, Mary

Available from Sheffield Hallam University Research Archive (SHURA) at:

http://shura.shu.ac.uk/32905/

This document is the author deposited version. You are advised to consult the publisher's version if you wish to cite from it.

#### **Published version**

DIB, Sarah, FAIR, Frankie, MCCANN, Lucy Jane, NICHOLLS, Antonia, KALEA, Anastasia Z, SOLTANI, Hora and FEWTRELL, Mary (2023). Effects of exclusive breastfeeding promotion interventions on child outcomes: a systematic review and meta-analysis. Annals of Nutrition and Metabolism.

#### Copyright and re-use policy

See http://shura.shu.ac.uk/information.html



## Annals of Nutrition and Metabolism

Ann Nutr Metab , DOI: 10.1159/000535564 Received: August 25, 2023 Accepted: November 25, 2023 Published online: December 5, 2023

# Effects of exclusive breastfeeding promotion interventions on child outcomes: a systematic review and meta-analysis

Dib S, Fair FJ, McCann LJ, Nicholls A, Kalea AZ, Soltani H, Fewtrell M

ISSN: 0250-6807 (Print), eISSN: 1421-9697 (Online) https://www.karger.com/ANM Annals of Nutrition and Metabolism

Disclaimer:

Accepted, unedited article not yet assigned to an issue. The statements, opinions and data contained in this publication are solely those of the individual authors and contributors and not of the publisher and the editor(s). The publisher and the editor(s) disclaim responsibility for any injury to persons or property resulting from any ideas, methods, instructions or products referred to the content.

#### Copyright:

This article is licensed under the Creative Commons Attribution 4.0 International License (CC BY) (http://www.karger.com/Services/OpenAccessLicense). Usage, derivative works and distribution are permitted provided that proper credit is given to the author and the original publisher.

© 2023 The Author(s). Published by S. Karger AG, Basel

#### Systematic Review & Meta-Analysis

## *Effects of exclusive breastfeeding promotion interventions on child outcomes: a systematic review and meta-analysis*

Sarah Dib<sup>a</sup>, Frankie Joy Fair<sup>b</sup>, Lucy Jane McCann<sup>c</sup>, Antonia Nicholls<sup>d</sup>, Anastasia Z. Kalea<sup>d,e</sup>, Hora Soltani<sup>b</sup>, Mary Fewtrell<sup>a</sup>

#### Affiliation:

<sup>a</sup> UCL Great Ormond Street Institute of Child Health, London, UK

<sup>b</sup> Health Research Institute, Sheffield Hallam University, Sheffield, UK

<sup>c</sup> Wolfson Institute of Population Health, Queen Mary University, London, UK

<sup>d</sup> Division of Medicine, University College London, London, UK

<sup>e</sup> Institute of Cardiovascular Science, University College London, UK

Short Title: Exclusive Breastfeeding Promotion & Child Outcomes Corresponding Author: Dr. Sarah Dib Population, Policy & Practice UCL Great Ormond Street Institute of Child Health 30 Guilford Street, London WC1N 1EH United Kingdom E-mail: <u>sarah.dib@ucl.ac.uk</u> Telephone: +447746771009

Number of Tables: 1 Number of Figures: 4 Word count: 3459 words Keywords: Exclusive Breastfeeding; Growth; Morbidity; Infection; Mortality

#### Abstract

**Introduction:** Interventions promoting exclusive breastfeeding (EBF) may benefit infant health outcomes, but evidence is inconsistent. The objective of this review was to assess the effect of interventions promoting EBF on health outcomes in infants and children under 7 years of age. **Methods:** A literature search was conducted using EMBASE, MEDLINE, CINAHL, Cochrane Central, Cochrane Database of Systematic Reviews, and WHO International Clinical Trials Registry Platform from inception to April 2022. Inclusion criteria were randomized or cluster-randomized controlled trials aiming to increase EBF that reported effects on offspring growth, morbidity and/or mortality up to age 7 years. The primary outcome was infant/child growth. Secondary outcomes were infant morbidity and mortality and exclusive breastfeeding rates. Data were pooled using a random-effects model. **Results:** 32 studies (40 papers) were identified. No effect on infant/child growth was observed. EBF promotion interventions significantly improved EBF rates up to 6 months (n=25; OR 3.15; 95%CI 2.36,4.19) and significantly reduced the odds of respiratory illness at 0-3 months by 59% (n=2; OR 0.41; 95%CI 0.20,0.84) but not at later time-points. A borderline significant effect was observed for diarrhea (n=12; OR 0.84; 95%CI 0.70,1.00). Effects on hospitalizations or mortality were not significant. **Discussion/Conclusion:** EBF promotion interventions improve EBF rates and might yield modest reductions in infant morbidity without affecting infant/child growth. Future studies should investigate the cost-effectiveness of these interventions and examine potential benefits on other health outcomes.

#### Introduction

Breast milk is the optimum source of nutrition for infants. Promoting and supporting breastfeeding is an important public health intervention with multiple benefits for infants and mothers, including a reduced risk of infant gastrointestinal and respiratory infections and reduced risk of breast cancer in the mother [1]. In 2001, following a systematic review on the effects on child and maternal health of exclusive breastfeeding (EBF) for six months versus EBF for three to four months [2] the World Health Organization (WHO) made a global recommendation that infants should be exclusively breastfed until six months of age. After six months, safe complementary foods are recommended with maintenance of breastfeeding up to two years of age or beyond. However, despite numerous initiatives over many years in different settings, globally only 44% of infants 0-6 months old are EBF [3], with particularly low rates in many Western countries.

Much of the available evidence focuses on health effects of breastfeeding rather than the impact of EBF or EBF for six months. Furthermore, defining causal effects of breastfeeding *per se*, or specific periods of EBF on health outcomes is problematic as most studies are observational, given the ethical and practical difficulties of conducting randomized trials in this field. Therefore, the optimal duration of EBF [4] and the magnitude of benefits in different settings remains uncertain. Observational studies attempt to control for possible confounding factors, but it is difficult to completely account for the complex biological, social, economic and cultural factors that influence breastfeeding and health outcomes.

Although the process of randomizing mothers to EBF for different durations is challenging, it is possible to randomize mothers to interventions aimed at promoting a longer duration of EBF. These trials provide an opportunity to systematically synthesize evidence for health outcomes in infants and children with different EBF exposure. A systematic review and meta-analysis examining the efficacy of interventions aimed at increasing exposure to EBF found that mothers who received interventions were 2.77 times more likely to EBF up to six months [5]. However, this analysis did not consider the impact on infant health outcomes.

To address this gap, we conducted a systematic review to evaluate the effect of interventions that aimed to increase infant exposure to EBF on their growth, morbidity and mortality up to seven years of age in both low- and high-income settings.

#### Methods

#### Search Strategy

The full protocol is available on PROSPERO [CRD42020203796]. A systematic search of the literature was performed using medical subject headings (MeSH) and key text words pertaining to EBF interventions and infant/child mortality and health outcomes (see <u>Supplementary 1</u>). The search was conducted in August 2020, and re-run on April 27th, 2022. The Cochrane Database of Systematic Reviews, Cochrane Central Register of Controlled Trails, WHO International Clinical Trials Registry Platform (ICTRP), Cumulative Index of Nursing and Allied Health Literature, EMBASE and MEDLINE were searched from inception. Reference lists of included articles were reviewed manually to identify additional sources. Studies were limited to those published in English and French.

#### **Study Inclusion and Exclusion Criteria**

Randomized controlled trials (RCTs) and cluster-randomized controlled trials (CRCTs) that focused on increasing infant exposure to EBF and reported at least one of the specified outcomes were included.

#### Population

Healthy mother-infant pairs and/or infants who were followed at any time up to age seven years were eligible. Studies that exclusively included preterm infants or mothers living with HIV were excluded.

#### Intervention

Any intervention(s) aiming to increase exposure to EBF.

#### Control

Controls could have no intervention or receive standard care for the study setting.

#### Outcomes

The primary outcome was infant and child growth, including anthropometric measurements such as weight, height/length, body composition and BMI at the latest reported time point and at 3-4 months, 6 months, and at ≥18 months, and changes in weight and length/height. Secondary outcomes were infant and child morbidity and mortality and EBF rates up to six months. Morbidity included prevalence or incidence of disease and/or infections at the latest

available time point and at 0-3 and 4-6 months. For all outcomes, we used author definitions but this was evaluated in the quality assessment.

#### Setting

The review included studies from low-middle income countries (LMIC) and high-income (HIC) countries, defined according the World Bank classification.

#### **Selection of Studies**

Two review authors independently ran the initial and updated search using Covidence (https://www.covidence.org/), scanning titles and abstracts of retrieved records; those meeting the selection criteria were assessed further through full text appraisal. Discrepancies were discussed with a third reviewer to reach consensus. Additionally, three other members of the review team checked a random sample of 10% of both included and excluded studies to ensure agreement with methodology and logic. Figure 1 shows the PRISMA [6] flowchart.

#### **Data Extraction and Management**

Data extraction was performed by two independent review authors. All data included in the meta-analysis were additionally checked by a third author. Key information was extracted using a pre-defined template on Covidence. Disagreements were resolved through discussion with the review team. For studies with incomplete data, the authors of the original study were contacted.

#### Assessment of Risk of Bias

Two authors independently assessed the risk of bias (ROB) for included studies using the Cochrane Handbook for Systematic Review of Interventions with disagreements resolved by discussion [7]. In addition, for CRCTs we assessed the risk of (i) recruitment bias; (ii) baseline imbalance; (iii) loss of clusters; (iv) incorrect analysis; and (v) comparability with individually-randomized trials as outlined in the Cochrane Handbook for Systematic Reviews of Interventions.[8] Each ROB category was classified 'high risk', 'low risk' or 'unclear risk'. As specified in the protocol, 'high quality' was defined as a trial having adequate sequence generation, allocation concealment and an attrition rate <25%.

#### **Data Synthesis**

For two or more studies reporting similar outcome measures, we undertook statistical analysis using the Review Manager software (RevMan 2014). We analyzed outcomes on an intention-to-treat basis whenever possible. Random-effects meta-analysis was used given variability in baseline characteristics, interventions and outcome reporting [9]. Results were presented as mean treatment effect with 95% confidence intervals (CI) using odds ratio for dichotomous data and standardized mean difference for continuous data.

For CRCTs that did not adjust for clustering in their analyses (or did not report adjusted results), we adjusted the sample sizes using the methods described in the Cochrane Handbook [Section 23.1][10] using an estimate of the intracluster correlation co-efficient (ICC) derived from the trial (if possible), from a similar trial or from a study in a similar population. Meta-analysis using the generic inverse variance method was undertaken in RevMan using the adjusted estimates from CRCT that adequately accounted for the cluster design within the analysis, alongside an approximated effect estimate for the CRCTs that did not.

Heterogeneity was assessed through the Chi<sup>2</sup> and I<sup>2</sup> statistics. An I<sup>2</sup> statistic of 50% or greater or Chi<sup>2</sup> significance level p<0.1 was regarded as indicating substantial heterogeneity between studies.[7]

#### Subgroup Analysis

To identify potential sources of substantial heterogeneity, subgroup analysis according to country income classification, type of intervention, and timing of intervention was performed when at least two studies were available in each comparator. We also planned subgroup analysis according to maternal BMI, Baby-Friendly Initiative accreditation and family socioeconomic status but data did not allow for these analyses. Additionally, subgroup analysis for child growth, morbidity and mortality outcomes was performed, stratified by whether the intervention resulted in a significant improvement in the odds of EBF at latest available timepoint, 6 months, 3-4 months, or 1-3 months.

#### **Sensitivity Analysis**

Sensitivity analysis was performed based on the quality of the included trials determined by ROB to identify the impact on the overall results. For cluster-randomized trials, sensitivity analysis was also used to investigate the effect of variation in the ICC and unit of randomization. We initially adjusted the sample size for unadjusted outcomes using the ICC of 0.1 [11] then undertook sensitivity analysis using varying ICC from 0.01 to 0.3 [12, 13].

#### **Assessment of Reporting Bias**

For outcomes that were reported in  $\geq$ 10 trials, we investigated reporting biases by visually examining asymmetry in funnel plots. We also undertook the Egger, Smith [14] regression asymmetry test using Stats Direct software, with p<0.10 taken as evidence of small study effects.

#### Results

#### **Studies Selected**

We initially identified 9,158 studies, of which 7,158 were screened and 106 full-text articles were reviewed (Fig. 1). A total of 32 studies (40 publications) were eligible for inclusion [11-13, 15-43].

#### **Study Characteristics**

Of the included studies (<u>Table 1</u>), 17 were RCTs and 15 were CRCTs, 28 were in LMIC and only 4 were in HIC, with sample size between 108 and 140 048 participants. Year of publication was 1994 to 2021, with 23 studies published in the last 10 years. Most interventions lasted longer than one month (25 interventions), four lasted one week to one month and four lasted one week or less (one study contained two interventions, one <one week and one >one month). The type of intervention varied, with most (n=20) involving peer counselling and/or home visits while five studies targeted hospital/clinic practices. Three studies investigated the effects of longer *vs* shorter durations of EBF by randomizing infants to complementary feeding from four months or EBF until six months. Two studies provided mobile phone-based support to breastfeeding mothers. Two studies involved breastfeeding support provided by community representatives or members. One study was undertaken in three separate countries; effect size estimates were given separately for each country, so they were entered into the meta-analyses separately.

#### **Quality Assessment**

The quality of included studies based on pre-defined criteria was poor, with 24/32 studies of low or uncertain quality and only eight of high quality (Supplementary 2).

Eighty-seven percent (14/16) of CRCTs had high risk of recruitment bias where individuals were recruited after randomization of clusters. All studies had low risk of bias when considering imbalance between randomized groups. Two studies had high ROB for not undertaking adjustment for clustering and two had potentially high ROB due to loss of clusters.

#### Meta-Analysis Results for Infant/Child Growth

EBF promotion interventions did not have an overall effect on infant weight (Fig. 2; SMD= -0.01 [95% CI -0.11, 0.09], 11 studies, 11,556 participants,  $I^2$  67%), infant length/height (Fig. 2; SMD = 0.01 [95% CI -0.03, 0.05], 12 studies, 28,817 participants,  $I^2$  21%) nor BMI/Weight-for-Length (Fig. 2; SMD = -0.04 [95% CI -0.10, 0.03], 11 studies, 27,702 participants,  $I^2$  54%) at the latest time-point reported. There were also no effects at 3-4 months, 6 months nor at ≥18 months (Supplementary 3). Similarly, there were no effects on gain in length or weight (Supplementary 3).

#### Meta-Analysis Results for Infant/Child Morbidity and Mortality

There was a trend towards reduced odds of infant diarrhea with EBF promotion interventions at the latest available time-point (Fig. 3; OR = 0.84 [95% CI 0.70, 1.00], 12 studies, 24,060 participants, I<sup>2</sup> 42%), but not at 0-3 months and 4-6 months (Supplementary 3). There was no effect of EBF promotion interventions on respiratory illness at the latest time-point (Fig. 3; OR = 0.80 [95% CI 0.60, 1.06], 5 studies, 19, 718 participants, I<sup>2</sup> 44%) nor at 4-6 months, however there was a significant reduction in the odds of respiratory illness at 0-3 months (Supplementary 3; OR = 0.41 [95%CI 0.20, 0.84], 2 studies, I<sup>2</sup> 0%). No differences were found in infant hospitalization (Fig. 3; OR = 0.56 [95% CI 0.31, 1.02], 5 studies, 3,162 participants, I<sup>2</sup> 53%) nor infant mortality (Fig. 3; OR = 0.98 [95% CI 0.75, 1.28], 5 studies, 60,918 participants, I<sup>2</sup> 52%).

#### **Meta-Analysis Results for EBF**

The interventions were successful at improving the odds of EBF (pooled odds ratio 3.15 [95% Cl 2.36, 4.19, 25 studies, 202,644 participants, l<sup>2</sup> 85%]) at any time point from 0-6 months (Fig. 4). Similar effects were observed at 6 months, 3-4 months, and 1-3 months separately (<u>Supplementary 3</u>). Across all settings, the median EBF rate at 3-4 months was 66.7% (IQR 25.4; 10.4-83.1) and 34.6% (IQR 26.7; 6.2-80.1) in the intervention and control groups (17 studies), respectively. This rate was lower at 6 months (34.5% (IQR 45; 2.0 -73.0) versus 13.7% (IQR 27.6; 0.0 -54.8)), respectively (18 studies).

#### Subgroup Analyses

(3.31 [2.44, 4.49] vs 1.66 [0.98, 2.81]; p=0.02) with 80% of the heterogeneity in EBF explained by income classification (Supplementary 4). There were significant subgroup differences in the OR of respiratory illness between interventions that started prenatally (0.51 [95% CI 0.33, 0.79]) vs postnatally (0.97 [95% CI 0.78, 1.20]) with 85% of heterogeneity explained by timing of the intervention. There were no significant differences in growth outcomes between subgroups stratified by whether the intervention resulted in an improvement in EBF (at latest available time point, 6 months, 3-4 months or 1-3 months). There were insufficient studies to test these subgroup differences for respiratory illness and mortality outcomes. Confining analyses to RCTs resulted in significantly lower pooled odds of diarrhea (OR = 0.69 [95% CI 0.53, 0.89]) and a trend towards higher length/height (SMD = 0.13 [95% CI 0.00, 0.26]) with interventions. Conversely, limiting

inclusion to high-quality studies did not notably affect the magnitude or direction of the intervention effect and confidence intervals. Using an ICC of 0.01 resulted in significantly lower pooled odds of diarrhea (OR = 0.83 [95% CI 0.70, 0.99]). However, overall there was minimal impact on the magnitude of the intervention effect and confidence intervals, and no change in the direction of effect for analyses using different ICC.

Studies conducted in LMIC yielded a significantly higher OR of EBF in the intervention group compared to those in HIC

#### **Publication Bias**

**Sensitivity Analysis** 

There was no evidence for publication bias for weight, length/height, BMI/weight-for-length, or diarrhea (Supplementary 5). There was evidence of publication bias (Egger bias = 1.81; 95% CI -0.09, 3.71; p=0.06 < 0.1) for EBF when using data for the latest available time point within each study, but this was not evident when including data at different timepoints (6 months, 3-4 months, or 1-3 months).

#### **Discussion/Conclusion**

This review examined the effects of exposure to more vs less breast milk as a result of EBF promotion interventions. The interventions significantly improved EBF rates at various time-points, however, EBF rates at 6 months were low overall. Promotion of EBF reduced the odds of respiratory illness at 0-3 months. There was also a trend towards a reduction in the overall odds of diarrhea which was significant in the sensitivity analysis. No significant differences were found in infant/child growth, even when results were stratified by whether the intervention improved EBF. There were also no significant differences in hospitalizations or mortality.

Our results for infant/child growth are consistent with a previous systematic review and meta-analysis published in 2012 that concluded that EBF for a longer duration was not associated with growth deficits [4], and add to the body of evidence that promotion of EBF does not give rise to growth concerns. Interestingly, our analyses confined to RCTs also suggested better linear growth in infants whose mothers received an EBF promotion intervention. Overall, there was a significant reduction in the odds of respiratory illness at 0-3 months following EBF interventions but not at 4-6 months or at the latest available time point. The inconsistent effects of EBF interventions on respiratory infection in different studies [11, 13, 16, 28, 31] are similar to those reported by Kramer and Kakuma [4] and contrast with previous studies that consistently show strong protective effects of breastfeeding per se against respiratory infection [44]. This might suggest that promotion of breastfeeding rather than focusing specifically on EBF could yield larger benefits in reducing respiratory infection, although this was not addressed in our meta-analysis. There was a trend towards lower odds of diarrhea with EBF promotion interventions. Contrary to what might be expected due to clustering 'herd effects', the effect on diarrhea was significant when considering only RCTs (and excluding CRCTs) and also when adjusting for clustering using a smaller ICC. Nevertheless, with the exception of three CRCTs (from Ethiopia [11], DR Congo [13], and South Africa [39, 45]) which showed higher prevalence/incidence of diarrhea in the intervention group, all other studies favored the intervention group. It is biologically plausible that promotion of EBF would protect against gastrointestinal infections as it avoids contamination from unsafe preparation of breast milk substitutes or other foods, and breast milk in general provides several anti-microbial and anti-inflammatory compounds. We were not able to explore this further as data on the infant feeding practices of

participants who were not EBF was not consistently reported.

Previous breastfeeding meta-analyses of mostly observational studies such as Victora et al. [1] found that breastfeeding offers strong protection against infections and hospitalizations due to infections. They also found that breastfeeding was associated with some reductions in overweight/obesity. Direct comparisons between these findings and ours are difficult as our focus was the promotion of EBF whereas Victora et al. [1] compared groups with

breastfeeding defined in several ways (for example, any breastfeeding vs none, predominant vs partial, EBF vs partial/predominant). We also only included interventional studies and excluded observational studies. Furthermore, most studies included in our meta-analysis reviewed the effects of exclusive EBF promotion interventions rather than the effects of EBF; only 3 trials directly randomized mothers to EBF for specific durations. Therefore, the magnitude of effects on infant and child outcomes might be underestimated by non-compliance with the interventions and depends on the success of the interventions in promoting EBF. It is also possible that had we been able to sub-divide the control groups according to breastfeeding intensity (predominant, vs partial vs not breastfeeding) or according to definition of not EBF (introduction of solids, liquids and/or infant formula), we may have seen larger effect sizes. Strengths of our review are the inclusion of studies from all settings and a range of health outcomes assessed up to seven years of age, allowing a comprehensive review of the effects of EBF promotion. However, there are several limitations. We examined the effects of exposure to more vs less breast milk as a result of EBF promotion interventions and did not directly address whether EBF for 6 months has greater benefits than EBF for 4 months; indeed, we noted that EBF rates at 6 months were low even in the intervention groups with a median of only 34.5% for studies that reported this outcome. As expected, included studies were heterogenous and country income classification, type of intervention and timing of the intervention were significant sources of heterogeneity for some outcomes. We were also unable to conduct planned subgroup analysis by maternal BMI, Baby-Friendly Initiative accreditation status, or family socioeconomic status. There were insufficient studies investigating other potential benefits of EBF promotion such as reduced risk of non-communicable diseases or allergy/asthma. Additionally, as for all meta-analyses, decisions must be made about which data to include, for example when multiple interventions are used or the same outcome is assessed in different ways and this can potentially influence findings. We used a systematic approach and discussed these decisions to avoid introducing bias.

In conclusion, EBF promotion interventions were successful at improving EBF. There were modest reductions in respiratory infection and diarrhea with no effects on infant and child growth or mortality. However, even modest decreases in infections, which are associated with significant morbidity and mortality especially in low-income settings, could translate to significant public health benefits and reduced healthcare expenses. Future studies investigating the effects of longer vs shorter EBF durations on other outcomes, and including cost-effectiveness analysis of EBF vs breastfeeding promotion could provide further insight into this issue.

#### **Statement of Ethics**

An ethics statement is not applicable because this study is based exclusively on published literature.

#### **Conflict of Interest Statement**

SD, FF, LM, AN, AZK, and HS have no conflicts of interest to declare. MF has received an unrestricted donation for research on infant nutrition from Philips (not related to the current manuscript); she is Assistant Officer for Nutrition at the Royal College of Paediatrics & Child Health UK, a member of the Infant Nutrition working group at the European Food Safety Authority (EFSA) and General Secretary of the European Society for Paediatric Gastroenterology, Hepatology & Nutrition (ESPGHAN).

#### **Funding Sources**

The authors received no financial support for this research.

#### **Author Contributions**

Dr Sarah Dib ran the final study search, assessed articles for inclusion, undertook quality appraisal, extraction of data, and data analysis, drafted the initial manuscript and critically reviewed and revised the manuscript. Dr Frankie Fair developed and reviewed the study protocol, checked a random sample of included and excluded studies, cross-checked all extracted data within the analyses and assisted with the data analysis, and critically reviewed and revised the manuscript. Dr Lucy McCann assessed articles for inclusion, cross-checked all extracted data, and critically reviewed and revised the manuscript. Dr Lucy McCann assessed articles for inclusion, cross-checked all extracted data, and critically reviewed and revised the manuscript. Ms Antonia Nicholls conceptualized the study, developed the study protocol, ran the initial search, assessed articles for inclusion in the initial search, and critically reviewed and revised the manuscript. Prof Anastasia Kalea developed the study protocol and critically reviewed and revised the manuscript. Prof Hora Soltani developed and reviewed the study protocol, checked a random sample of included and excluded studies, cross checked and critically reviewed and revised the manuscript. Prof Mary Fewtrell conceptualized the study, developed the study protocol, undertook quality appraisal of included studies, discussed discrepancies in study inclusion, checked a random sample of included and excluded studies, and critically reviewed and revised the manuscript. All authors approved the final manuscript as submitted and agree to be accountable for all aspects of the work.

#### **Data Availability Statement**

All data generated or analyzed during this study are included in this article and its online supplementary material files. Further inquiries can be directed to the corresponding author.

#### References

1. Victora CG, Bahl R, Barros AJ, França GV, Horton S, Krasevec J, et al. Breastfeeding in the 21st century: epidemiology, mechanisms, and lifelong effect. The lancet. 2016;387(10017):475-90.

2. Kramer MS, Kakuma R. Optimal duration of exclusive breastfeeding. Cochrane Database of Systematic Reviews. 2002(1).

3. World Health Organization. Infant and young child feeding. 2021. Available from: https://www.who.int/news-room/fact-sheets/detail/infant-and-young-child-feeding.

4. Kramer MS, Kakuma R. Optimal duration of exclusive breastfeeding. Cochrane database of systematic reviews. 2012(8).

5. Kim SK, Park S, Oh J, Kim J, Ahn S. Interventions promoting exclusive breastfeeding up to six months after birth: A systematic review and meta-analysis of randomized controlled trials. International Journal of Nursing Studies. 2018;80:94-105.

6. Page MJ, McKenzie JE, Bossuyt PM, Boutron I, Hoffmann TC, Mulrow CD, et al. The PRISMA 2020 statement: an updated guideline for reporting systematic reviews. International journal of surgery. 2021;88:105906.

7. Higgins JP, Green S. Cochrane handbook for systematic reviews of interventions. The Cochrane Collaboration. London, UK. 2011.

8. Higgins JP, Altman DG, Gøtzsche PC, Jüni P, Moher D, Oxman AD, et al. The Cochrane Collaboration's tool for assessing risk of bias in randomised trials. Bmj. 2011;343.

9. Wood L, Egger M, Gluud LL, Schulz KF, Jüni P, Altman DG, et al. Empirical evidence of bias in treatment effect estimates in controlled trials with different interventions and outcomes: meta-epidemiological study. bmj. 2008;336(7644):601-5.

Downloaded from http://karger.com/anm/article-pdf/doi/10.1159/000535564/4054122/000535564.pdf by guest on 20 December 2023

10. Higgins Jpt T, Chandler J, Cumpston M, Li T, Page M, Welch V. Cochrane Handbook for Systematic Reviews of Interventions version 6.3 (updated February 2022). Cochrane; 2022.

11. Abdulahi M, Fretheim A, Argaw A, Magnus JH. Breastfeeding education and support to improve early initiation and exclusive breastfeeding practices and infant growth: a cluster randomized controlled trial from a rural Ethiopian setting. Nutrients. 2021;13(4):1204.

12. Ara G, Khanam M, Papri N, Nahar B, Kabir I, Sanin KI, et al. Peer counseling promotes appropriate infant feeding practices and improves infant growth and development in an urban slum in Bangladesh: a community-based cluster randomized controlled trial. Current Developments in Nutrition. 2019;3(7):nzz072.

13. Yotebieng M, Labbok M, Soeters HM, Chalachala JL, Lapika B, Vitta BS, et al. Ten Steps to Successful Breastfeeding programme to promote early initiation and exclusive breastfeeding in DR Congo: a cluster-randomised controlled trial. The Lancet Global Health. 2015;3(9):e546-e55.

14. Egger M, Smith GD, Schneider M, Minder C. Bias in meta-analysis detected by a simple, graphical test. Bmj. 1997;315(7109):629-34.

15. Agudelo SI, Gamboa OA, Acuña E, Aguirre L, Bastidas S, Guijarro J, et al. Randomized clinical trial of the effect of the onset time of skin-to-skin contact at birth, immediate compared to early, on the duration of breastfeeding in full term newborns. International Breastfeeding Journal. 2021;16(1):1-10.

16. Bashour HN, Kharouf MH, AbdulSalam AA, El Asmar K, Tabbaa MA, Cheikha SA. Effect of postnatal home visits on maternal/infant outcomes in Syria: a randomized controlled trial. Public Health Nursing. 2008;25(2):115-25.

17. Bhandari N, Bahl R, Mazumdar S, Martines J, Black RE, Bhan MK. Effect of community-based promotion of exclusive breastfeeding on diarrhoeal illness and growth: a cluster randomised controlled trial. The lancet. 2003;361(9367):1418-23.

18. Birungi N, Fadnes LT, Okullo I, Kasangaki A, Nankabirwa V, Ndeezi G, et al. Effect of breastfeeding promotion on early childhood caries and breastfeeding duration among 5 year old children in eastern Uganda: a cluster randomized trial. PloS one. 2015;10(5):e0125352.

19. Chapman DJ, Morel K, Anderson AK, Damio G, Pérez-Escamilla R. Breastfeeding peer counseling: from efficacy through scale-up. Journal of Human Lactation. 2010;26(3):314-26.

20. Cohen RJ, Brown KH, Dewey K, Canahuati J, Rivera LL. Effects of age of introduction of complementary foods on infant breast milk intake, total energy intake, and growth: a randomised intervention study in Honduras. The Lancet. 1994;344(8918):288-93.

21. Cui R, Wang E. The effect of postpartum family visits on the promotion of breastfeeding and improvement of maternal and infant health. American Journal of Translational Research. 2021;13(12):14089.

22. Davies-Adetugbo AA, Adetugbo K, Orewole Y, Fabiyi A. Breast-feeding promotion in a diarrhoea programme in rural communities. Journal of Diarrhoeal Diseases Research. 1997:161-6.

23. Dewey KG, Cohen RJ, Brown KH, Rivera LL. Age of introduction of complementary foods and growth of term, low-birth-weight, breast-fed infants: a randomized intervention study in Honduras. The American journal of clinical nutrition. 1999;69(4):679-86.

24. Fadnes LT, Nankabirwa V, Engebretsen IM, Sommerfelt H, Birungi N, Lombard C, et al. Effects of an exclusive breastfeeding intervention for six months on growth patterns of 4–5 year old children in Uganda: the cluster-randomised PROMISE EBF trial. BMC Public Health. 2016;16(1):1-9.

25. Fang Y, Zhu L, Bao L. The effect of multi-dimensional postpartum visits on increasing the breastfeeding rate of parturients with inverted nipple: a randomised study. Annals of Palliative Medicine. 2021;10(3):3078-85.

26. Gabida M, Chemhuru M, Tshimanga M, Gombe NT, Takundwa L, Bangure D. Effect of distribution of educational material to mothers on duration and severity of diarrhoea and pneumonia, Midlands Province, Zimbabwe: a cluster randomized controlled trial. International breastfeeding journal. 2015;10(1):1-12.

27. Hanson C, Manzi F, Mkumbo E, Shirima K, Penfold S, Hill Z, et al. Effectiveness of a home-based counselling strategy on neonatal care and survival: a cluster-randomised trial in six districts of rural southern Tanzania. PLoS medicine. 2015;12(9):e1001881.

28. Hmone MP, Li M, Agho K, Dibley M, editors. Impact of SMS text messages to improve exclusive breastfeeding and reduce other adverse infant feeding practices in Yangon, Myanmar: a randomized controlled trial. Annals of Nutrition and Metabolism; 2017: KARGER ALLSCHWILERSTRASSE 10, CH-4009 BASEL, SWITZERLAND.

29. Jakobsen MS, Sodemann M, Biai S, Nielsen J, Aaby P. Promotion of exclusive breastfeeding is not likely to be cost effective in West Africa. A randomized intervention study from Guinea-Bissau. Acta Paediatrica. 2008;97(1):68-75.

30. Khan AI, Hawkesworth S, Ekström EC, Arifeen S, Moore SE, Frongillo EA, et al. Effects of exclusive breastfeeding intervention on child growth and body composition: the MINIM at trial, B angladesh. Acta Paediatrica. 2013;102(8):815-23.

31. Kramer MS, Chalmers B, Hodnett ED, Sevkovskaya Z, Dzikovich I, Shapiro S, et al. Promotion of Breastfeeding Intervention Trial (PROBIT): a randomized trial in the Republic of Belarus. Jama. 2001;285(4):413-20.

32. le Roux IM, Tomlinson M, Harwood JM, O'Connor MJ, Worthman CM, Mbewu N, et al. Outcomes of home visits for pregnant mothers and their infants: a cluster randomized controlled trial. Aids. 2013;27(9):1461-71.

33. Lewycka S, Mwansambo C, Rosato M, Kazembe P, Phiri T, Mganga A, et al. Effect of women's groups and volunteer peer counselling on rates of mortality, morbidity, and health behaviours in mothers and children in rural Malawi (MaiMwana): a factorial, cluster-randomised controlled trial. The Lancet. 2013;381(9879):1721-35.

34. Morandi A, Tommasi M, Soffiati F, Destro F, Fontana L, Grando F, et al. Prevention of obesity in toddlers (PROBIT): a randomised clinical trial of responsive feeding promotion from birth to 24 months. International Journal of Obesity. 2019;43(10):1961-6.

35. Morrow AL, Guerrero ML, Shults J, Calva JJ, Lutter C, Bravo J, et al. Efficacy of home-based peer counselling to promote exclusive breastfeeding: a randomised controlled trial. The Lancet. 1999;353(9160):1226-31.

36. Nair N, Tripathy P, Sachdev H, Pradhan H, Bhattacharyya S, Gope R, et al. Effect of participatory women's groups and counselling through home visits on children's linear growth in rural eastern India (CARING trial): a cluster-randomised controlled trial. The Lancet Global Health. 2017;5(10):e1004-e16.

37. Nikièma L, Huybregts L, Martin-Prevel Y, Donnen P, Lanou H, Grosemans J, et al. Effectiveness of facilitybased personalized maternal nutrition counseling in improving child growth and morbidity up to 18 months: A cluster-randomized controlled trial in rural Burkina Faso. PloS one. 2017;12(5):e0177839.

38. Ogaji DS, Arthur AO, George I. Effectiveness of mobile phone-based support on exclusive breastfeeding and infant growth in nigeria: a randomized controlled trial. Journal of Tropical Pediatrics. 2021;67(1):fmaa076.

39. Tylleskär T, Jackson D, Meda N, Engebretsen IMS, Chopra M, Diallo AH, et al. Exclusive breastfeeding promotion by peer counsellors in sub-Saharan Africa (PROMISE-EBF): a cluster-randomised trial. The Lancet. 2011;378(9789):420-7.

40. Schwartz R, Vigo Á, Dias de Oliveira L, Justo Giugliani ER. The effect of a pro-breastfeeding and healthy complementary feeding intervention targeting adolescent mothers and grandmothers on growth and prevalence of overweight of preschool children. PLoS One. 2015;10(7):e0131884.

41. Sehhatie FS, Mirghafourvand M, Havizari S. Effect of prenatal counseling on exclusive breastfeeding frequency and infant weight gain in mothers with previous unsuccessful breastfeeding: a randomized controlled clinical trial. The Journal of Maternal-Fetal & Neonatal Medicine. 2020;33(21):3571-8.

42. Wells JC, Jonsdottir OH, Hibberd PL, Fewtrell MS, Thorsdottir I, Eaton S, et al. Randomized controlled trial of 4 compared with 6 mo of exclusive breastfeeding in Iceland: differences in breast-milk intake by stable-isotope probe. The American journal of clinical nutrition. 2012;96(1):73-9.

43. Wen LM, Baur LA, Simpson JM, Rissel C, Flood VM. Effectiveness of an early intervention on infant feeding practices and "tummy time": a randomized controlled trial. Archives of pediatrics & adolescent medicine. 2011;165(8):701-7.

44. Horta BL, Victora CG, World Health Organization. Short-term effects of breastfeeding: a systematic review on the benefits of breastfeeding on diarrhoea and pneumonia mortality. 2013.

45. Engebretsen IMS, Nankabirwa V, Doherty T, Diallo AH, Nankunda J, Fadnes LT, et al. Early infant feeding practices in three African countries: the PROMISE-EBF trial promoting exclusive breastfeeding by peer counsellors. International breastfeeding journal. 2014;9(1):1-11.

#### **Figure Legends**

Fig. 1. PRISMA Flow Chart

Fig. 2. Forest plot of intervention vs control comparison for outcomes: a) infant weight, b) infant length/height, and c) infant BMI/weight-for-length at 3->18 months.

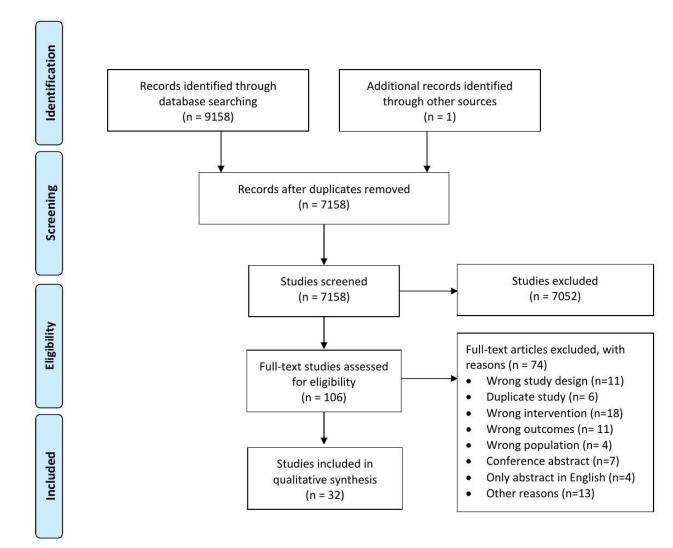
Contributing studies are sorted in chronological order. Square data markers represent effect size estimates (SMD), with size of the markers corresponding to 95% CIs and diamond data markers representing the overall effect size based on included studies.

Fig. 3. Forest plot of intervention vs control comparison for outcomes: a) infant diarrhea, b) respiratory illness, c) hospitalization and d) mortality.

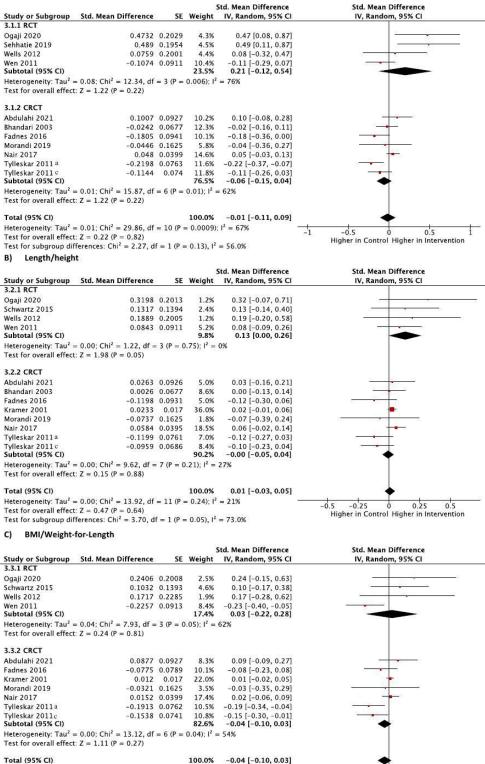
Contributing studies are sorted in chronological order. Square data markers represent effect size estimates (OR), with size of the markers corresponding to 95% CIs and diamond data markers representing the overall effect size based on included studies.

Fig. 4. Forest plot of intervention vs control comparison for exclusive breastfeeding

Contributing studies are sorted in chronological order. Square data markers represent effect size estimates (OR), with size of the markers corresponding to 95% CIs and diamond data markers representing the overall effect size based on included studies.

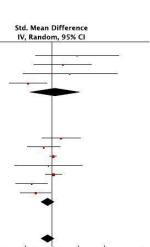


#### A) Weight



Heterogeneity: Tau<sup>2</sup> = 0.00; Chi<sup>2</sup> = 21.68, df = 10 (P = 0.02); I<sup>2</sup> = 54% Test for overall effect: Z = 1.15 (P = 0.25) Test for subgroup differences: Chi<sup>2</sup> = 0.26, df = 1 (P = 0.61), l<sup>2</sup> = 0%

Downloaded from http://karger.com/anm/article-pdf/doi/10.1159/000535564/4054122/000535564.pdf by guest on 20 December 2023



#### A) Diarrhea

Study or Subgroup 4.1.1 RCT	log[Odds Ratio	SE	Weight	Odds Ratio IV, Random, 95% C	Odds Ratio IV, Random, 95% Cl
4.1.1 KCI					
Bashour 2008	-0.31	4 0.1443	15.1%	0.73 [0.55, 0.97	n
Davies-Adetugbo 1997		2 0.4346		0.51 [0.22, 1.19	
Hmone 2017		2 0.4852		0.50 [0.19, 1.29	
Subtotal (95% CI)	-0.057	2 0.4052	21.9%	0.69 [0.53, 0.89	
Heterogeneity: Tau <sup>2</sup> =	0.00 Chi <sup>2</sup> - 1.11	4f _ 2 (P _		25 25	
Test for overall effect: .			- 0.37), 1	- 0/6	
4.1.2 CRCT					
Abdulahi 2021	0.311	9 0.3324	5.7%	1.37 [0.71, 2.62	1
Bhandari 2003	-0.168	4 0.0784	20.8%	0.85 [0.72, 0.99	)]
Gabida 2015	-0.313	5 0.5616	2.3%	0.73 [0.24, 2.20	
Kramer 2001	-0.505	8 0.2094	10.6%	0.60 [0.40, 0.91	.1
Morrow 1999	-0.924	3 0.4963	2.9%	0.40 [0.15, 1.05	5]
Tylleskar 2011a	0.307	5 0.2233	9.8%	1.36 [0.88, 2.11	1
Tylleskar 2011b		7 0.3474		0.81 [0.41, 1.60	
Tylleskar 2011c		2 0.2113		0.81 [0.53, 1.22	
Yotebieng 2015	0.246	9 0.2186		1.28 [0.83, 1.96	
Subtotal (95% CI)			78.1%	0.90 [0.73, 1.11	J 🕈
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: .			= 0.06); 1	<sup>2</sup> = 46%	
Total (95% CI)			100.0%	0.84 [0.70, 1.00	n 📥
Heterogeneity: Tau <sup>2</sup> =	0.03 Chi <sup>2</sup> = 18.96	df = 11 (			
Test for overall effect: 1			- 0.00),	1 - 72/0	0.2 0.5 1 2
Test for subgroup diffe			P = 0.11	$l^2 = 60.0\%$	Higher in Control Higher in Interve
B) Respiratory Illne		$y_{1} = 1 (1)$	- 0.11),		
	:33			Odds Ratio	Odds Ratio
Study or Subgroup 4.2.1 RCT	log[Odds Ratio]	SE V	Weight IV	, Random, 95% Cl	IV, Random, 95% Cl
Bashour 2008	0.025	0 1442	32.2%	1.03 [0.77, 1.36]	
Basnour 2008 Hmone 2017	-0.8267				· · · · · · · · · · · · · · · · · · ·
Subtotal (95% CI)	-0.8267	0.5705	11.3% 43.5%	0.44 [0.21, 0.92] 0.72 [0.32, 1.64]	
	0.38. Chi2 4.40	df _ 1 /2			
Heterogeneity: Tau <sup>2</sup> = Test for overall effect:			- 0.03); l'	- / 0/0	
4.2.2 CRCT					
Abdulahi 2021	-0.5816	0.2783	17.2%	0.56 [0.32, 0.96]	A
Kramer 2001	-0.1393	0.1982	25.1%	0.87 [0.59, 1.28]	· · · · · · · · · · · · · · · · · · ·
Yotebieng 2015	-0.0619		14.2%	0.94 [0.50, 1.77]	
Subtotal (95% CI)			56.5%	0.78 [0.59, 1.05]	•
Heterogeneity: Tau <sup>2</sup> =			= 0.36); 1	2 = 3%	
Test for overall effect:	Z = 1.65 (P = 0.10)	)			
			0.13)	0.80 [0.60, 1.06]	<b>_</b>
사망을 같은 것은 것이라. 상황한 가지만 An and and a set of the set of	0.04. Chi2 7	ar = 4 (P)	= 0.13); 1	= 44%	0.1 0.2 0.5 1 2 5
Heterogeneity: Tau <sup>2</sup> =		)			
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe	Z = 1.58 (P = 0.11)		(P = 0.85)	$1^2 = 0\%$	Higher in Control Higher in Interven
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe C) Hospitalization	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0	4, df = 1		Odds Ratio	Odds Ratio
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe C) Hospitalization Study or Subgroup	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 log[Odds Ratio]	4, df = 1 SE \	Weight IN	Odds Ratio /, Random, 95% Cl	
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe C) Hospitalization Study or Subgroup Agudelo 2021	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 log[Odds Ratio] -0.5242	4, df = 1 <u>SE V</u> 0.5288	Weight IN 18.5%	Odds Ratio /, Random, 95% Cl 0.59 [0.21, 1.67]	Odds Ratio
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe C) Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 log[Odds Ratio] -0.5242 -0.3967	4, df = 1 <u>SE V</u> 0.5288 0.3194	Weight IN 18.5% 28.3%	Odds Ratio /, Random, 95% Cl 0.59 [0.21, 1.67] 0.67 [0.36, 1.26]	Odds Ratio
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe ) Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 log[Odds Ratio] -0.5242 -0.3967 -1.1705	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292	Weight IN 18.5% 28.3% 18.5%	Odds Ratio /, Random, 95% CI 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88]	Odds Ratio
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe <b>() Hospitalization</b> Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 log[Odds Ratio] -0.5242 -0.3967 -1.1705 -2.5946	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633	Weight IX 18.5% 28.3% 18.5% 6.9%	Odds Ratio /, Random, 95% Cl 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60]	Odds Ratio
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe The spitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 log[Odds Ratio] -0.5242 -0.3967 -1.1705	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633	Weight IN 18.5% 28.3% 18.5%	Odds Ratio /, Random, 95% CI 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88]	Odds Ratio
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe ) Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 log[Odds Ratio] -0.5242 -0.3967 -1.1705 -2.5946	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633 0.3295	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         27.8%	Odds Ratio (, Random, 95% CI 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05]	Odds Ratio
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe C) Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% CI)	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 -0.5242 -0.3967 -1.1705 -2.5946 0.0735	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633 0.3295	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         100.0%	Odds Ratio 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02]	Odds Ratio
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe <b>C) Hospitalization</b> Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% CI) Heterogeneity: Tau <sup>2</sup> =	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 <u>log[Odds Ratio]</u> -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48,	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633 0.3295 df = 4 (P	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         100.0%	Odds Ratio 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02]	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe C) Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% CI) Heterogeneity: Tau <sup>2</sup> =	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 <u>log[Odds Ratio]</u> -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48,	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633 0.3295 df = 4 (P	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         100.0%	Odds Ratio (, Random, 95% Cl 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02]	Odds Ratio IV, Random, 95% CI
Study or Subgroup Agudelo 2021	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 <u>log[Odds Ratio]</u> -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48,	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633 0.3295 df = 4 (P	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         100.0%	Odds Ratio 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] <sup>2</sup> = 53%	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe C) Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 <u>log[Odds Ratio]</u> -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06	4, df = 1 <u>SE (</u> ) 0.5288 0.3194 0.5292 1.0633 0.3295 J df = 4 (P )	Weight         IX           18.5%         28.3%           18.5%         6.9%           27.8%         27.8%           100.0%         = 0.08); I <sup>+</sup>	Odds Ratio 4, Random, 95% Cl 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] 2 = 53% Odds Ratio	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe <b>C)</b> Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 <u>log[Odds Ratio]</u> -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48,	4, df = 1 <u>SE (</u> ) 0.5288 0.3194 0.5292 1.0633 0.3295 J df = 4 (P )	Weight         IX           18.5%         28.3%           18.5%         6.9%           27.8%         27.8%           100.0%         = 0.08); I <sup>+</sup>	Odds Ratio 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] <sup>2</sup> = 53%	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe ) Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% Cl) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality Study or Subgroup 4.4.1 RCT	Z = 1.58 (P = 0.11 rences: Chi <sup>2</sup> = 0.0 log[Odds Ratio] -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06 log[Odds Ratio]	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633 0.3295 J df = 4 (P ) <u>SE V</u>	Weight         IN           18.5%         28.3%           6.9%         27.8%           100.0%         = 0.08); I <sup>+</sup> Weight         IN	Odds Ratio 4, Random, 95% CI 0.59 (0.21, 1.67) 0.67 (0.36, 1.26) 0.31 (0.11, 0.88) 0.07 (0.01, 0.60) 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] 2 = 53% Odds Ratio 4, Random, 95% CI	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe ) Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality Study or Subgroup 4.4.1 RCT Bashour 2008	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 <u>log[Odds Ratio]</u> -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06	4, df = 1 <u>SE (1</u> ) 0.5288 0.3194 0.5292 1.0633 0.3295 J df = 4 (P ) <u>SE (1</u> )	Weight         IN           18.5%         28.3%           6.9%         27.8%           100.0%         = 0.08); I <sup>+</sup> Weight         IN	Odds Ratio 4, Random, 95% Cl 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] 2 = 53% Odds Ratio	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe ) Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality Study or Subgroup 4.4.1 RCT Bashour 2008 Jakobsen 2008	Z = 1.58 (P = 0.11 rences: Chi <sup>2</sup> = 0.0 <u>log[Odds Ratio]</u> -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06 <u>log[Odds Ratio]</u> 1.7399	4, df = 1 <u>SE (1</u> ) 0.5288 0.3194 0.5292 1.0633 0.3295 J df = 4 (P ) <u>SE (1</u> )	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         27.8%           100.0%         0.08); I'           Weight         IN           0.8%         5	Odds Ratio 4, Random, 95% CI 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] 2 = 53% Odds Ratio 4, Random, 95% CI 5.70 [0.31, 103.37]	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe ) Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% Cl) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality Study or Subgroup 4.4.1 RCT Bashour 2008 Jakobsen 2008 Subtotal (95% Cl)	Z = 1.58 (P = 0.11 erences: Chi <sup>2</sup> = 0.0 log[Odds Ratio] -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06 log[Odds Ratio] 1.7399 0.6019	4, df = 1 <u>SE 1</u> 0.5288 0.3194 0.5292 1.0633 0.3295 J df = 4 (P ) <u>SE 1</u> 1.4788 0.4414	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         27.8%           100.0%         = 0.08); I'           Weight         IN           0.8%         5           7.9%         8.7%	Odds Ratio /, Random, 95% CI 0.59 (0.21, 1.67) 0.67 (0.36, 1.26) 0.31 (0.11, 0.88) 0.07 (0.01, 0.60) 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] 2 = 53% Odds Ratio /, Random, 95% CI 1.83 (0.77, 4.34) 1.83 (0.77, 4.34) 2.00 [0.87, 4.59]	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe ) Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality Study or Subgroup 4.4.1 RCT Bashour 2008 Jakobsen 2008 Subtotal (95% CI) Heterogeneity: Tau <sup>2</sup> =	Z = 1.58 (P = 0.11 rences: Chi <sup>2</sup> = 0.0 -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06 log[Odds Ratio] 1.7399 0.6019 0.00; Chi <sup>2</sup> = 0.54,	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633 0.3295 df = 4 (P ) <u>SE V</u> 1.4788 0.4414 df = 1 (P	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         27.8%           100.0%         = 0.08); I'           Weight         IN           0.8%         5           7.9%         8.7%	Odds Ratio /, Random, 95% CI 0.59 (0.21, 1.67) 0.67 (0.36, 1.26) 0.31 (0.11, 0.88) 0.07 (0.01, 0.60) 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] 2 = 53% Odds Ratio /, Random, 95% CI 1.83 (0.77, 4.34) 1.83 (0.77, 4.34) 2.00 [0.87, 4.59]	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe (2) Hospitalization Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% Cl) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality Study or Subgroup 4.4.1 RCT	Z = 1.58 (P = 0.11 rences: Chi <sup>2</sup> = 0.0 -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06 log[Odds Ratio] 1.7399 0.6019 0.00; Chi <sup>2</sup> = 0.54, Z = 1.64 (P = 0.10	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633 0.3295 df = 4 (P ) <u>SE V</u> 1.4788 0.4414 df = 1 (P )	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         27.8%           100.0%         = 0.08); I'           Weight         IN           0.8%         5           7.9%         8.7%	Odds Ratio 4, Random, 95% CI 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] 2 = 53% Odds Ratio 4, Random, 95% CI 5.70 [0.31, 103.37] 1.83 [0.77, 4.34] 2.00 [0.87, 4.59] 2 = 0%	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffect ) Hospitalization Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality Study or Subgroup 4.4.1 RCT Bashour 2008 Subtotal (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: 4.4.2 CRCT Hanson 2015	Z = 1.58 (P = 0.11 rences: Chi <sup>2</sup> = 0.0 10g[Odds Ratio] -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06 10g[Odds Ratio] 1.7399 0.6019 0.00; Chi <sup>2</sup> = 0.54, Z = 1.64 (P = 0.10 0.0766	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633 0.3295 df = 4 (P ) <u>SE V</u> 1.4788 0.4414 df = 1 (P ) 0.0707	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         100.0%           100.0%         0.08); I'           0.8%         5           7.9%         8.7%           8.7%         9.46); I'           42.6%         42.6%	Odds Ratio /, Random, 95% CI 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] <sup>2</sup> = 53% Odds Ratio /, Random, 95% CI 5.70 [0.31, 103.37] 1.83 [0.77, 4.34] 2.00 [0.87, 4.59] <sup>2</sup> = 0% 1.08 [0.94, 1.24]	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffect: Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% Cl) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality Study or Subgroup 4.4.1 RCT Bashour 2008 Subtotal (95% Cl) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: 4.4.2 CRCT Hanson 2015 Lewycka 2013	Z = 1.58 (P = 0.11 rences: Chi <sup>2</sup> = 0.0 log[Odds Ratio] -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06 log[Odds Ratio] 1.7399 0.6019 0.00; Chi <sup>2</sup> = 0.54, Z = 1.64 (P = 0.10 0.07666 -0.091	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633 0.3295 df = 4 (P ) <u>SE V</u> 1.4788 0.4414 df = 1 (P ) 0.0707 0.1392	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         27.8%           100.0%         0.08); I'           0.8%         5           7.9%         8.7%           42.6%         32.0%	Odds Ratio /, Random, 95% CI 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] <sup>2</sup> = 53% Odds Ratio /, Random, 95% CI 1.83 [0.77, 4.34] 2.00 [0.87, 4.59] <sup>2</sup> = 0%	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffect C) Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality Study or Subgroup 4.4.1 RCT Bashour 2008 Jakobsen 2008 Jakobsen 2008 Jakobsen 2008 Subtotal (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: 4.4.2 CRCT Hanson 2015 Lewycka 2013 Nair 2017	Z = 1.58 (P = 0.11 rences: Chi <sup>2</sup> = 0.0 10g[Odds Ratio] -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06 10g[Odds Ratio] 1.7399 0.6019 0.00; Chi <sup>2</sup> = 0.54, Z = 1.64 (P = 0.10 0.0766	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633 0.3295 df = 4 (P ) <u>SE V</u> 1.4788 0.4414 df = 1 (P ) 0.0707 0.1392	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         100.0%           100.0%         = 0.08); 1°           0.8%         5           7.9%         8.7%           8.7%         32.0%           16.6%         16.6%	Odds Ratio /, Random, 95% CI 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.01 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] 2 = 53% Odds Ratio /, Random, 95% CI 5.70 [0.31, 103,37] 1.83 [0.77, 4.34] 2.00 [0.87, 4.59] 2 = 0% 1.08 [0.94, 1.24] 0.91 [0.70, 1.20] 0.60 [0.35, 1.01]	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffect ) Hospitalization Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality Study or Subgroup 4.4.1 RCT Bashour 2008 Subtotal (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: 4.4.2 CRCT Hanson 2015 Lewycka 2013 Nair 2017 Subtotal (95% CI)	Z = 1.58 (P = 0.11 rences: Chi <sup>2</sup> = 0.0 10g[Odds Ratio] -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06 10g[Odds Ratio] 1.7399 0.6019 0.00; Chi <sup>2</sup> = 0.54, Z = 1.64 (P = 0.10 0.0766 -0.091 -0.5158	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633 0.3295 df = 4 (P ) <u>SE V</u> 1.4788 0.4414 df = 1 (P ) 0.0707 0.1392 0.2695	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         100.0%           20.08); I'         0.08); I'           0.8%         5           7.9%         8.7%           8.7%         9.046); I'           42.6%         32.0%           16.6%         91.3%	Odds Ratio /, Random, 95% CI 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] <sup>2</sup> = 53% Odds Ratio /, Random, 95% CI 5.70 [0.31, 103.37] 1.83 [0.77, 4.34] 2.00 [0.87, 4.59] <sup>2</sup> = 0% 1.08 [0.94, 1.24] 0.91 [0.70, 1.20] 0.60 [0.37, 1.19]	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffect ) Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality Study or Subgroup 4.4.1 RCT Bashour 2008 Jakobsen 2008 Subtotal (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: 4.4.2 CRCT Hanson 2015 Lewycka 2013 Nair 2017	Z = 1.58 (P = 0.11 rences: Chi <sup>2</sup> = 0.0 <b>log[Odds Ratio]</b> -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06 <b>log[Odds Ratio]</b> 1.7399 0.6019 0.00; Chi <sup>2</sup> = 0.54, Z = 1.64 (P = 0.10 0.0766 -0.091 -0.5158 0.03; Chi <sup>2</sup> = 5.22,	4, df = 1 <u>SE 1</u> 0.5288 0.3194 0.5292 1.0633 0.3295 df = 4 (P ) <u>SE 1</u> 1.4788 0.4414 df = 1 (P ) 0.0707 0.1392 0.2695 df = 2 (P	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         100.0%           20.08); I'         0.08); I'           0.8%         5           7.9%         8.7%           8.7%         9.046); I'           42.6%         32.0%           16.6%         91.3%	Odds Ratio /, Random, 95% CI 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] <sup>2</sup> = 53% Odds Ratio /, Random, 95% CI 5.70 [0.31, 103.37] 1.83 [0.77, 4.34] 2.00 [0.87, 4.59] <sup>2</sup> = 0% 1.08 [0.94, 1.24] 0.91 [0.70, 1.20] 0.60 [0.37, 1.19]	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe J Hospitalization Study or Subgroup Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D Mortality Study or Subgroup 4.4.1 RCT Bashour 2008 Subtotal (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: 4.4.2 CRCT Hanson 2015 Lewycka 2013 Nair 2017 Subtotal (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Subtotal (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Subtotal (95% CI) Heterogeneity: Tau <sup>2</sup> = Test for overall effect:	Z = 1.58 (P = 0.11 rences: Chi <sup>2</sup> = 0.0 <b>log[Odds Ratio]</b> -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06 <b>log[Odds Ratio]</b> 1.7399 0.6019 0.00; Chi <sup>2</sup> = 0.54, Z = 1.64 (P = 0.10 0.0766 -0.091 -0.5158 0.03; Chi <sup>2</sup> = 5.22,	4, df = 1 <u>SE V</u> 0.5288 0.3194 0.5292 1.0633 0.3295 df = 4 (P ) <u>SE V</u> 1.4788 0.4414 df = 1 (P ) 0.0707 0.1392 0.2695 df = 2 (P )	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         100.0%           100.0%         0.08); I'           0.8%         5.7%           42.6%         32.0%           16.6%         91.3%           = 0.07); I'         10.07); I'	Odds Ratio /, Random, 95% CI 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.31 [0.11, 0.88] 0.07 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] <sup>2</sup> = 53% Odds Ratio /, Random, 95% CI 5.70 [0.31, 103.37] 1.83 [0.77, 4.34] 2.00 [0.87, 4.59] <sup>2</sup> = 0% 1.08 [0.94, 1.24] 0.91 [0.70, 1.20] 0.60 [0.35, 1.01] 0.92 [0.71, 1.19] <sup>2</sup> = 62%	Odds Ratio IV, Random, 95% CI
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Test for subgroup diffe ) Hospitalization Agudelo 2021 Bashour 2008 Chapman 2013 Fang 2021 Jakobsen 2008 Total (95% Cl) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: D) Mortality Study or Subgroup 4.4.1 RCT Bashour 2008 Jakobsen 2008 Subtotal (95% Cl) Heterogeneity: Tau <sup>2</sup> = Test for overall effect: 4.4.2 CRCT Hanson 2015 Lewycka 2013 Nair 2017 Subtotal (95% Cl)	Z = 1.58 (P = 0.11 rences: Chi <sup>2</sup> = 0.0 <b>log[Odds Ratio]</b> -0.5242 -0.3967 -1.1705 -2.5946 0.0735 0.23; Chi <sup>2</sup> = 8.48, Z = 1.90 (P = 0.06 <b>log[Odds Ratio]</b> 1.7399 0.6019 0.00; Chi <sup>2</sup> = 0.54, Z = 1.64 (P = 0.10 0.0766 -0.091 -0.5158 0.03; Chi <sup>2</sup> = 5.22, Z = 0.62 (P = 0.54)	4, df = 1 SE V 0.5288 0.3194 0.5292 1.0633 0.3295 df = 4 (P ) SE V 1.4788 0.4414 df = 1 (P ) 0.0707 0.1392 0.2695 df = 2 (P )	Weight         IN           18.5%         28.3%           18.5%         6.9%           27.8%         100.0%           100.0%         = 0.08); I'           0.8%         5           7.9%         8.7%           8.7%         = 0.46); I'           42.6%         32.0%           16.6%         91.3%           = 0.07); I'         100.0%	Odds Ratio /, Random, 95% Cl 0.59 [0.21, 1.67] 0.67 [0.36, 1.26] 0.7 [0.01, 0.60] 1.08 [0.56, 2.05] 0.56 [0.31, 1.02] 2 = 53% Odds Ratio /, Random, 95% Cl 3:70 [0.31, 103.37] 1.83 [0.77, 4.34] 2:00 [0.87, 4.59] 2 = 0% 1.08 [0.94, 1.24] 0.91 [0.70, 1.20] 0.92 [0.71, 1.19] 2 = 62% 0.98 [0.75, 1.28]	IV, Random, 95% CI

500 B B B		1000		Odds Ratio	Odds Ratio
Study or Subgroup	log[Odds Ratio]	SE	Weight	IV, Random, 95% CI	IV, Random, 95% Cl
5.1.1 RCT					
Agudelo 2021	0.0898	0.2688	4.8%	1.09 [0.65, 1.85]	
Bashour 2008	0.4966	0.1838	5.2%	1.64 [1.15, 2.36]	
Chapman 2013	2.0307	1.523	0.8%	7.62 [0.39, 150.77]	s
Davies-Adetugbo 1997	2.3523	0.479	3.5%	10.51 [4.11, 26.87]	
Fang 2021	1.0917	0.3914	4.0%	2.98 [1.38, 6.42]	
Hmone 2017	1.4418	0.2911	4.6%	4.23 [2.39, 7.48]	
Khan 2013	0.967	0.1346	5.4%	2.63 [2.02, 3.42]	
Ogaji 2020	0.5182	0.3532	4.3%	1.68 [0.84, 3.36]	
Sehhatie 2019	2.5649	0.475	3.6%	13.00 [5.12, 32.98]	
Wen 2011	0.7337	0.5073	3.4%	2.08 [0.77, 5.63]	
Subtotal (95% CI)			39.6%	2.93 [1.92, 4.45]	•
Heterogeneity: Tau <sup>2</sup> = 0.	.30; Chi <sup>2</sup> = 41.15, di	f = 9 (P	< 0.0000	1); $I^2 = 78\%$	HERBORN D
Test for overall effect: Z	= 5.01 (P < 0.0000)	1)			
5.1.2 CRCT					
Abdulahi 2021	0.6098	0.1932	5.2%	1.84 [1.26, 2.69]	
Ara 2019	1.5053	0.2887	4.6%	4.51 [2.56, 7.93]	
Bhandari 2003	2.8679	0.5033	3.4%	17.60 [6.56, 47.20]	· · · · · · · · · · · · · · · · · · ·
Gabida 2015	0.3379	0.4937	3.4%	1.40 [0.53, 3.69]	
Hanson 2015	0.7885	0.1024	5.5%	2.20 [1.80, 2.69]	200
Kramer 2001	2.6615	1.1141	1.3%	14.32 [1.61, 127.12]	
Le Roux 2013	1.2782	0.3221	4.5%	3.59 [1.91, 6.75]	
Lewycka 2013	1.3137	0.3564	4.2%	3.72 [1.85, 7.48]	
Morandi 2019	0.3496	0.3226	4.4%	1.42 [0.75, 2.67]	
Morrow 1999	2.2877	0.5736	3.0%	9.85 [3.20, 30.32]	· · · · ·
Nair 2017	0.077	0.1793	5.2%	1.08 [0.76, 1.53]	
Tylleskar 2011 a	2.0528	0.1811	5.2%	7.79 [5.46, 11.11]	
Tylleskar 2011b	1.7138	0.7665	2.2%	5.55 [1.24, 24.93]	· · · · · · · · · · · · · · · · · · ·
Tylleskar 2011c	2.2435		3.2%	9.43 [3.23, 27.47]	
Yotebieng 2015	0.1906		4.9%	1.21 [0.75, 1.94]	
Subtotal (95% CI)			60.4%	3.34 [2.22, 5.03]	•
Heterogeneity: $Tau^2 = 0$ .	.49; Chi <sup>2</sup> = 115.74, 6	df = 14	P < 0.00	001); $I^2 = 88\%$	
Test for overall effect: Z			88 - 23 SAN	nar Sanal Alba	
Total (95% CI)			100.0%	3.15 [2.36, 4.19]	•
Heterogeneity: $Tau^2 = 0$ .	.38: Chi <sup>2</sup> = 156.96.	df = 24	P < 0.00	$001$ ); $I^2 = 85\%$	
Test for overall effect: Z					0.01 0.1 1 10 1
Test for subgroup differe			- 0.66)	$1^2 - 0\%$	Higher in Control Higher in Intervention

#### Table 1. Overview of Included Studies

Author Name and Date	Country	Study Design	Population	Sample Size	Intervention Description	Duration of Intervention	Breastfeeding Outcomes	Infant/Child Outcomes
Abdulahi 2021[11]	Ethiopia (LMIC)	CRCT	Pregnant women (2nd/3rd trimester)	468	BF education and support intervention delivered by peer-supporters	7 months starting at the 8th month of pregnancy	↑ EBF at 6 months; ↑ EBI	↔ WAZ, LAZ, underweight, stunting, wasting at 6 months; ↓ acute respiratory illness at 6 months
Agudelo 2021[15]	Colombia (LMIC)	RCT	Mothers of healthy FT infants	297	Immediate skin-to-skin contact (in the first minute after birth)	60 minutes starting post- delivery	$\leftrightarrow$ EBF at 3 and 6 months	$\leftrightarrow$ %weight change between birth and the first week of life
Ara 2019[12]	Bangladesh (LMIC)	CRCT	Married pregnant women with <4 living children	350	IYCF counselling and psychosocial stimulation education provided by peer counsellors involving the mothers and key family members	Before delivery until infant was 11 months old (~11-19 months)	$\Lambda$ EBF at 1, 3, and 5 months	↑ Change in length over 12 months
Bashour 2008[16]	Syria (LMIC)	RCT	Mothers of healthy FT infants	876	Home visits conducted by midwives (Group A: 4 HV on days 1, 3, 7 and 30 post-delivery; Group B: 1 HV on day 3)	Varies but a maximum of 30 days in group A	↑ EBF at 4 months in Groups A and B; ↔ BF	↔ Diarrhea, jaundice, fever, cold, cough or infection at 4 months
Bhandari 2003[17]	India (LMIC)	CRCT	Mother-infant pairs	1115	Health and nutrition workers were trained to counsel mothers for EBF and to deliver messages promoting EBF to community representatives	24 months starting at birth	↑ EBF at 3, 4, 5 and 6 months	$\leftrightarrow$ Mean weight, length, WAZ <2, LAZ <2 at 3 and 6 months; $\downarrow$ Diarrhea (in previous 7 days) at 3 and 6 months
Birungi 2015[18, 39]	Uganda (LMIC)	CRCT	Pregnant women intending to BF	765	EBF peer-counselling during pregnancy and postpartum	~29 weeks starting in the third trimester	$\leftrightarrow$ BF duration at 5 years	$\leftrightarrow$ Dental caries at 5 years
Chapman 2013[46]	USA (HIC)	RCT	Participants considering BF and had a pre-pregnancy BMI >27 and income <185% of the federal poverty level	154	Routine care plus prenatal, hospital and postpartum visits from a specialized BF peer counsellor	Varies but at least 10 weeks starting prenatally	↔ EBF and BF at 2 weeks, 1 month, 3 months and 6 months	$\leftrightarrow$ Otitis media and emergency department visits at 3 and 6 months; $\downarrow$ hospitalization at 3 and 6 months; $\uparrow$ diarrhea at 6 months
Cohen 1994[20]	Honduras (LMIC)	RCT	First-time healthy low- income mothers of FT infants weighing at least 2kg willing to EBF for 26 weeks	141	SF: Introduction of CF at 4 months with ad libitum nursing from 4-6 months; SF-M: introduction of CF at 4 months with maintenance of baseline nursing frequency from 4-6 months	2 months starting at 4 months		$\leftrightarrow$ Weight gain from 16-26 weeks; $\leftrightarrow$ diarrhea at 26 weeks; $\uparrow$ coughs (and respiratory illness p=0.05)
Cui and Wang 2021[21]	China (LMIC)	RCT	Mothers of healthy FT infants who were between 25-35 years	200	Routine care plus postpartum family visits. Mothers were also given a diet plan and encouraged to eat more protein.	Does not specify	$\uparrow$ EBF at 1, 2 and 3 months.	$\leftrightarrow$ Length and weight at 3 months; $\downarrow$ Incidence of adverse events (red buttocks, eczema, jaundice, umbilical infection) at 3 months
Davies- Adetugbo 1997[22]	Nigeria (LMIC)	RCT	Mothers who took their infants (<=3 months of age) to a primary care facility for uncomplicated acute diarrhea	161	3 sessions of BF counselling and lactation management at days 0, 2 and 7 to solve BF issues faced and to promote EBF	7 days, starting at day infants were presented to the primary care facility (Mean age: Control mean=55.6 (22.0); Intervention mean= 52.9 (21.4) days).	↑ EBF at days 7 and 21	$\leftrightarrow$ Recurrence of diarrhea by day 21
Dewey 1999[23]	Honduras (LMIC)	RCT	Mothers of FT infants weighing 1.5-2.5 kg at birth, who were willing to EBF for 6 mo	119	Mothers were advised to initiate complementary feeding at 4 months while maintaining baseline (at 4 months) BF frequency.	2 months starting at 4 months of age	$\downarrow$ Breast milk intake at 6 months	$\leftrightarrow$ Weight, length, head circumference change from 16 to 26 weeks, WAZ & LAZ in first 12 months; $\leftrightarrow$ %days with fever or respiratory illness from 16 to 26 weeks; $\uparrow$ %days with diarrhea from 16 to 26 weeks
Fadnes 2016[24]	Uganda (LMIC)	CRCT	Pregnant women intending to BF	765	EBF peer-counselling during pregnancy and postpartum	~29 weeks starting in the third trimester		$\uparrow$ Stunting at 2 years and underweight at 5 years
Fang 2021[25]	China (LMIC)	RCT	Mothers of singleton FT infants, who have inverted nipples and successfully breastfed at the hospital.	114	Nulti-dimensional postpartum visits involving online support, continuing health education in the community, and home visits	Does not specify	$\uparrow$ EBF at 1, 3 and 6 months	$\psi$ Incidence of infant hospitalization at 6 months
Gabida 2015[26]	Zimbabwe (LMIC)	CRCT	Mothers in antenatal care register who delivered within the selected clusters	357	Routine care plus cIYCF training for village health workers in two groups and provision of a BF newsletter in two groups. The newsletter contained non-financial incentives to encourage mothers to EBF until at least 14 weeks.	One time newsletter at delivery	↑ EBF at 14 weeks in the newsletter group; ↔ EBF at 20 weeks and in cIYCF group	↓ Recurrent episodes of diarrhea at 20 weeks in newsletter groups; ↓ Pneumonia in newsletter groups and cIYCF groups at 20 weeks; ↔ Morbidity at 14 weeks
Hanson 2015[27]	Tanzania (LMIC)	CRCT	All pregnant women in intervention wards (groups of 3-4 villages)	140048	Home-based counselling intervention on issues including hygiene, EBF, and care for LBW babies	From as soon as pregnancy identified until early postpartum.	$\uparrow$ EBF for the first 3 days; $\uparrow$ EBI	$\leftrightarrow$ All-cause neonatal mortality rate in the first 28 d of life

Hmone 2017[28]	Myanmar (LMIC)	RCT	Women from 28 to 34 weeks gestation who could access a networked mobile phone and had an uncomplicated singleton pregnancy	353	BF promotional text messages were sent 3 times per week in the evening	Over 9 months: from the time of recruitment until 6 months postdelivery.	↑ EBF over 6 months; ↑ BF over 6 months; ↔ EBI	$\downarrow$ ALRI at 3 and 5 months and over 6 months; $\downarrow$ Diarrhea at 3 months and over 6 months; $\leftrightarrow$ ALRI at 1 month, Diarrhea at 1 and 5 months, Fever or cold at 1, 3 and 5 months or over 6 months
Jakobsen 2008[29]	Guinea-Bissau (LMIC)	RCT	Mothers living in the area during pregnancy and present when visited by the field assistant	1721	Home visits involving education focused on encouraging mothers to postpone introduction of water and weaning food until the age of 4-6 months	Varies but from birth to 6 months of age unless the infant was reported to have started both water and weaning food.	Weaning food was significantly delayed	$\downarrow$ Weight at 4-6 months; $\leftrightarrow$ Weight at 7-12 days; $\leftrightarrow$ Diarrhea or hospitalizations in first 6 months; $\leftrightarrow$ Mortality in first 6 months
Khan 2013[30]	Bangladesh (LMIC)	RCT	Pregnant women (30 weeks) who were previously participating in the MINIMat trial	2845	EBF counselling provided by trained counsellors on a one-to-one basis (but could also include key family members) at home over 8 visits (2- last trimester, 1-within 7 days of delivery, 5- monthly intervals up to 6 months)	8 months starting at 8th month of pregnancy	$\uparrow$ EBF at 4 and 6 months	$\leftrightarrow$ Child growth from birth to 54 months of age (WHZ, HAZ, and WAZ)
Kramer 2001[31, 47, 48]	Belarus (LMIC)	CRCT	Healthy mothers who intended to BF and their healthy, FT singleton infants	17046	The experimental intervention was modelled on the BFHI and included intervention polyclinics to provide postnatal support	Does not specify	个 EBF at 3 and 6 months; 个 BF at 3, 6,9, and 12 months	$\leftrightarrow$ Height, BMI, waist and hip circumferences, triceps and subscapular skinfold thickness at 6.5 years; $\downarrow$ GI infections and atopic eczema in first 12 months; $\leftrightarrow$ Respiratory tract infections in first 12 months, systolic and diastolic blood pressure at 6.5 years, infant allergy and asthma at 6.5 years; $\leftrightarrow$ Infant mortality
Le Roux 2013[32, 49]	South Africa (LMIC)	CRCT	Pregnant women	1238	PIP: Standard care plus home visits by community healthcare workers where messages were provided on good maternal nutrition and preparing for BF; regular antenatal clinic attendance; HIV testing and prevention, stopping alcohol, BF and growth monitoring; medical adherence (immunizations, prevention for HIV-exposed children); infant bonding; and securing the child grant.	Started antenatally and up to 12 months post birth	↑ EBF at 6 months	个 HAZ > -2 at 6 months and WAZ > -2 at 18 months; ↔ WAZ <-2, WHZ <-2 at 6 months
Lewycka 2013[33]	Malawi (LMIC)	CRCT	All women aged 10-49 living in the study area who consented to participate	26262	Women's group intervention where women's groups were established supported by a cluster facilitator who was trained to discuss and help with maternal and child health problems. Volunteer peer counselling intervention delivered by trained counsellors who made five home visits during and after pregnancy (third trimester, within 1st week after birth, 1 month, 3 months and 5 months) Combination of WGI and VPC	Does not specify	↑ EBF at 6 months	↔IMR
Morandi 2019[34]	Italy (HIC)	CRCT	Primary pediatricians of healthy FT infants	569	Pediatricians were trained to provide parents with information about BF, feeding on demand, responsive feeding, timely CF, and other obesity prevention behaviors at all routine visits scheduled at 1, 3, 6, 12 and 24 months	23 months starting at 1 month	$\leftrightarrow$ EBF at 3 months; $\leftrightarrow$ BF at 3 and 6 months	<ul> <li>↑ Length at 3 months; ↔ Weight, W/L or BMI at 3,</li> <li>6, 12 and 24 months or Length at 6, 12 or 24 months;</li> <li>↔ Overweight/Obesity at 24 months</li> </ul>
Morrow 1999[35]	Mexico (LMIC)	CRCT	All pregnant women identified by a semiannual door-to-door census.	130	In the 3-visit group, peer-counsellors visited in late pregnancy and in the first and second weeks postpartum. In the six-visit group, peer- counsellors also visited in mid-pregnancy and at weeks 4 and 8 postpartum. These visits encouraged and helped with BF and EBF, and included key family members.	~21- 35 weeks for the 6 visit group starting at mid- pregnancy and ~14 weeks for the 3 visit group starting at late-pregnancy	↑ EBF at 3 months in 3-visit and 6-visit groups; ↑ EBF duration in 3-visit and 6-visit groups	$\mathbf \psi$ Diarrhea in the first 3 months (cumulative incidence)

Nair 2017[36]	India (LMIC)	CRCT	Pregnant women identified in rural districts in eastern India	5781	Community based workers conducted a single home visit to each pregnant woman in the third trimester of pregnancy for counselling on maternal nutrition followed by monthly home visits with counselling for growth promotion and IYCF. They also facilitated 2-3 participatory meetings with local women's groups per month to address underlying causes of undernutrition including birth spacing, nutrition in pregnancy, water, sanitation, and women's agency.	Nearly 2 years starting at last trimester of pregnancy until 2 years of age	$\leftrightarrow$ EBF at 6 months	↓ Underweight; $\leftrightarrow$ LAZ, WAZ, MUAC, stunting, wasting at 18 months; $\leftrightarrow$ Diarrhea, cough or fever at 6 months; ↓ Infant mortality
Nikiema 2017[37]	Burkina Faso (LMIC)	CRCT	Pregnant women in the catchment areas of the 12 selected health centers	2253	Nutrition counselling intervention that was implemented within the usual care environment, where the provider's training was focused on pregnant women's diet, BF and CF	>18 months starting during pregnancy	↑ EBF at 6 months; $\leftrightarrow$ EBI	$\leftrightarrow$ Wasting and stunting; $\leftrightarrow$ Incidence of diarrhea, fever or ARI; $\uparrow$ Mean incidence of child illness
Ogaji 2020[38]	Nigeria (LMIC)	RCT	Mothers of healthy infants who initiated BF after delivery in baby-friendly hospital	150	Mobile phone-based support plus usual care, where the women received monthly advisory support service from the same pediatrician. An average of 8 phone calls were made during which the mothers were reminded of the benefits of EBF and questions related to BF and the wellbeing of the mother and baby were answered.	6 months starting at 1 week post-delivery	$\leftrightarrow$ EBF at 6 months	↑ Weight, Length, WAZ at 6 months and weight gain over 6 months
Schwartz 2015[40]	Brazil (LMIC)	RCT	Adolescent mothers (<20 years) who live in the same household as their own mothers recruited from a baby-friendly hospital	323	Sessions at the maternity ward and at the participants' homes at 7, 15, 30, 60 and 120 days post-delivery during which advice on EBF, infant feeding challenges and complementary feeding and supporting material were given	120 days starting at birth	↑ EBF duration; ↔ BF durations	$\leftrightarrow$ BMI for age, HAZ, overweight (%), obesity (%) and stunting (%) at 4-7 years.
Sehhatie 2019[41]	Iran (LMIC)	RCT	Pregnant women (third trimester) who visited the health-care centers and had an unsuccessful previous BF experience and a singleton pregnancy	108	BF counselling sessions in groups of 5-7, four counselling sessions were held with a one-week interval during the third trimester. Phone or if necessary in-person counselling was offered to mothers on day 15, 2 months and the end of the month 4 postpartum.	Varies, around 1-5 months	个 EBF at 15 days, 2 months and 4 months	个Weight at 15 days; ↔ Weight at 2 months and 4 months
Tylleskar 2011- a[39, 45, 50]	Burkina Faso (LMIC)	CRCT	Pregnant women intending to BF	794	EBF peer-counselling during pregnancy and postpartum	~29 weeks starting in the third trimester	↑ EBF at 3 and 6 months; $↔$ EBI	$\downarrow$ WLZ at 12 and 24 weeks; $\uparrow$ Wasting at 12 weeks; $\leftrightarrow$ Diarrhea prevalence at 12 and 24 weeks
Tylleskar 2011-b[39, 45, 50]	Uganda (LMIC)	CRCT	Pregnant women intending to BF	765	EBF peer-counselling during pregnancy and postpartum	~29 weeks starting in the third trimester	个 EBF at 3 and 6 months; 个 EBI	$\downarrow$ WLZ at 24 weeks; $\downarrow$ WAZ at 12 and 24 weeks; $\uparrow$ Wasting at 12 and 24 weeks; $\leftrightarrow$ Diarrhea prevalence at 12 and 24 weeks
Tylleskar 2011- c[39, 45, 50]	South Africa (LMIC)	CRCT	Pregnant women intending to BF	1020	EBF peer-counselling during pregnancy and postpartum	~29 weeks starting in the third trimester	↑ EBF at 3 and 6 months; ↔ EBI	$\uparrow$ WLZ at 24 weeks; $\leftrightarrow$ Diarrhea prevalence at 12 and 24 weeks
Wells 2012[42, 51]	Iceland (HIC)	RCT	Mothers of healthy, FT, EBF infants at well-baby clinics	119	Mothers were asked to continue EBF until 6 months of age	2 months starting at 4 months of infant's age	↑Breast milk intake	↔Lean mass, fat mass, WAZ, LAZ, HAZ or BMI-for- age at various time points from 6-38 months
Wen 2011[43, 52]	Australia (HIC)	RCT	Pregnant women (24-34 gestation) attending antenatal clinics	667	The trained community nurse visited families 8 times at home, once at 30-36 weeks gestation and seven times after the birth (at 1, 3, 5, 9, 12, 18 and 24 months) where she taught the mother specific skills and knowledge in relation to healthy infant feeding practices and active play.	At least 28 months starting 30-36 weeks of pregnancy	$\leftrightarrow$ EBF at 6 months; $\uparrow$ BF at 6 and 12 months	↓BMI at 24 months

Yotebieng	
2015[13,	DR Congo
53]	(LMIC)

Women who had a healthy singleton birth in the randomized health facilities and who intended to attend well-baby clinic visits

CRCT

975[13]

931[53]

BFHI steps 1-9 were implemented at the selected facilities or BFHI steps 1-10 where steps 1-9 were implemented at the facilities, support was provided in well-child clinics and flyers were distributed to address the main BF barriers.

Hospital stay for those at steps 1-9 and 24 weeks for those at steps 1-10.

↑ EBF at 1, 6, and 14 weeks in Steps 1-9 and Steps 1-10 groups; ↑ EBF at 24 weeks in Steps 1-9 group;  $\leftrightarrow$  EBF at 24 weeks in Steps 1-10 group;  $\leftrightarrow$  EBI

↓ Diarrhea prevalence at 24 weeks in Steps 1-9 group; ↑Diarrhea prevalence at 14 and 24 weeks in Steps 1-10 group; ↔ Respiratory infection prevalence at 14 and 24 weeks in both groups

BF: Breastfeeding; BFHI: Baby-Friendly Health Initiative; CF: Complementary Feeding; cIYCF: Community Infant and Young Child Feeding; CRCT: Cluster-Randomized Controlled Trial; EBF: Exclusive Breastfeeding; FT: Full-Term; HAZ: Height-for-Age Z-score; HIC: High-Income Country; HV: Home Visit; LAZ: Length-for-Age Z-score; LMIC: Low- and Middle-Income Country; RCT: Randomized Controlled Trial; WAZ: Weight-for-Age Z-score; WLZ: Weight-for-Length Z-score.