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# On optimal tests for circular reflective symmetry about an unknown central direction

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## Abstract

Symmetry is one of the most fundamental of dividing hypotheses, its rejection, or not, heavily influencing subsequent modeling strategies. In this paper, the authors construct tests for circular reflective symmetry about an unknown central direction that are asymptotically valid within a semi-parametric class of distributions and maintain certain parametric local and asymptotic optimality properties. The asymptotic distributions of the test statistics under the null hypothesis and under local alternatives are established, and a pre-existing omnibus test is identified as a special case of the proposed construction. The finite-sample properties of the semi-parametric tests are compared with those of other testing approaches in a simulation experiment, and recommendations made regarding testing for reflective symmetry in practice. Analyses of data on the directions of cracks in hip replacements illustrate the proposed methodology.

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## References

Abe T, Pewsey A (2011) Sine-skewed circular distributions. *Stat Pap* 52:683–707

[Article](#) [MathSciNet](#) [Google Scholar](#)

Abe T, Kubota Y, Shimatani K, Aakala T, Kuuluvainen T (2012) Circular distributions of fallen logs as an indicator of forest disturbance regimes. *Ecol Indic* 18:559–566

[Article](#) [Google Scholar](#)

Azzalini A, Capitanio A (2003) Distributions generated by perturbation of symmetry with emphasis on a multivariate skew- $t$  distribution. *J R Stat Soc Ser B* 65:367–389

[Article](#) [MathSciNet](#) [Google Scholar](#)

Bogdan M, Bogdan K, Futschik A (2002) A data driven smooth test for circular uniformity. *Ann Inst Stat Math* 54:29–44

Jammalamadaka SR, SenGupta A (2001) Topics in circular statistics. World Scientific, Singapore

[Book](#) [Google Scholar](#)

Jones MC, Pewsey A (2005) A family of symmetric distributions on the circle. *J Am Stat Assoc* 100:1422–1428

[Article](#) [MathSciNet](#) [Google Scholar](#)

Jones MC, Pewsey A (2012) Inverse Batschelet distributions for circular data. *Biometrics* 68:183–193

[Article](#) [MathSciNet](#) [Google Scholar](#)

Jupp PE, Spurr B (1983) Sobolev tests for symmetry of directional data. *Ann Stat* 11:1225–1231

[Article](#) [MathSciNet](#) [Google Scholar](#)

Jupp PE, Regoli G, Azzalini A (2016) A general setting for symmetric distributions and their relationship to general distributions. *J Multivar Anal* 148:107–119

[Article](#) [MathSciNet](#) [Google Scholar](#)

Kato S, Jones MC (2015) A tractable and interpretable four-parameter family of unimodal distributions on the circle. *Biometrika* 102:181–190

[Article](#) [MathSciNet](#) [Google Scholar](#)

Kreiss J (1987) On adaptive estimation in stationary ARMA processes. *Ann Stat* 15:112–133

[Article](#) [MathSciNet](#) [Google Scholar](#)

Le Cam L, Yang G (2000) Asymptotics in statistics. Some basic concepts, 2nd edn. Springer, New York

Ley C, Verdebout T (2014) Simple optimal tests for circular reflective symmetry about a specified median direction. *Stat Sin* 24:1319–1339

[MathSciNet](#) [MATH](#) [Google Scholar](#)

Mann KA, Gupta S, Race A, Miller MA, Cleary RJ (2003) Application of circular statistics in the study of crack distribution around cemented femoral components. *J Biomech* 36:1231–1234

[Article](#) [Google Scholar](#)

Meintanis S, Verdebout T (2018) Le Cam maximin tests for symmetry of circular data based on the characteristic function. *Stat Sin* 29:1301–1320

[MathSciNet](#) [MATH](#) [Google Scholar](#)

Oliveira M, Crujeiras RM, Rodríguez-Casal A (2012) A plug-in rule for bandwidth selection in circular density estimation. *Comput Stat Data Anal* 56:3898–3908

[Article](#) [MathSciNet](#) [Google Scholar](#)

Pérez IA, Sánchez ML, García MA, Pardo N (2012) Analysis of  $\text{CO}_2$  daily cycle in the low atmosphere at a rural site. *Sci Total Environ* 431:286–292

[Article](#) [Google Scholar](#)

Pewsey A (2002) Testing circular symmetry. *Can J Stat* 30:591–600

[Article](#) [MathSciNet](#) [Google Scholar](#)

Pewsey A (2004) Testing for circular reflective symmetry about a known median axis. *J Appl Stat* 31:575–585

[Article](#) [MathSciNet](#) [Google Scholar](#)

Umbach D, Jammalamadaka SR (2009) Building asymmetry into circular distributions. *Stat Probab Lett* 79:659–663

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## Author information

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### Authors and Affiliations

Statistics Section, Department of Mathematics, KU Leuven, Leuven, Belgium

Jose Ameijeiras-Alonso

Department of Applied Mathematics, Computer Science and Statistics, Ghent University, Ghent, Belgium

Christophe Ley

Department of Mathematics, University of Extremadura, Cáceres, Spain

Arthur Pewsey

Mathematics Department, Université Libre de Bruxelles, Brussels, Belgium

Thomas Verdebout

### Corresponding author

Correspondence to [Jose Ameijeiras-Alonso](#).

### Additional information

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## Electronic supplementary material

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Below is the link to the electronic supplementary material.

[Supplementary material 1 \(pdf 138 KB\)](#)

## Appendices

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### Proof of Lemma 1

We show that  $\hat{\Gamma}_{f_0, g_0; 11} - \Gamma_{f_0, g_0; 11} = o_P(1)$  as  $n \rightarrow \infty$  under  $\mathbf{P}_{(\mu, 0)'; g_0}^{(n)}$ . Showing that  $\hat{\Gamma}_{g_0, p; 12} - \Gamma_{g_0, p; 12} = o_P(1)$  as  $n \rightarrow \infty$  under  $\mathbf{P}_{(\mu, 0)'; g_0}^{(n)}$  proceeds along the same lines. In this proof, we set  $\mu^{(n)} := \mu + n^{-1/2} \tau_1^{(n)}$  for some bounded sequence  $\tau_1^{(n)}$  as in Theorem 1. Because of the local discreteness of  $\hat{\mu}^{(n)}$  (Assumption B), it is sufficient to show that

$$n^{-1} \sum_{i=1}^n \dot{\varphi}_{f_0}(\Theta_i - \mu^{(n)}) - \mathbf{E}_{g_0}[\dot{\varphi}_{f_0}(\Theta_i - \mu)] = o_P(1)$$

as  $n \rightarrow \infty$  under  $\mathbf{P}_{(\mu, 0)'; g_0}^{(n)}$ . The law of large numbers leads to

$$n^{-1} \sum_{i=1}^n \dot{\varphi}_{f_0}(\Theta_i - \mu) - \mathbf{E}_{g_0}[\dot{\varphi}_{f_0}(\Theta_i - \mu)] = o_P(1)$$

as  $n \rightarrow \infty$  under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$  so that it only remains to show that

$$S_n := n^{-1} \sum_{i=1}^n \left( \dot{\varphi}_{f_0} \left( \Theta_i - \mu^{(n)} \right) - \dot{\varphi}_{f_0} \left( \Theta_i - \mu \right) \right)$$

is  $o_{\mathbf{P}}(1)$  as  $n \rightarrow \infty$  under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$ . As the  $\Theta_i$  are i.i.d., we find

$$\begin{aligned} \mathbf{E}[|S_n|] &\leq n^{-1} \sum_{i=1}^n \mathbf{E}[|\dot{\varphi}_{f_0} \left( \Theta_i - \mu^{(n)} \right) - \dot{\varphi}_{f_0} \left( \Theta_i - \mu \right)|] \\ &= \mathbf{E}[|\dot{\varphi}_{f_0} \left( \Theta_1 - \mu^{(n)} \right) - \dot{\varphi}_{f_0} \left( \Theta_1 - \mu \right)|]. \end{aligned}$$

Since  $\dot{\varphi}_{f_0}$  is continuous on a compact support, it is bounded. The result then follows by applying Lebesgue's dominated convergence theorem.  $\square$

## Proof of Lemma 2

To prove (i), we start by showing that  $\Delta_{f_0,j}^{(n)\star}(\hat{\mu}^{(n)}) - \Delta_{f_0,j}^{(n)\star}(\mu) = o_{\mathbf{P}}(1)$  under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$  as  $n \rightarrow \infty$ . First note that, due to Assumption B, we have under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$  as  $n \rightarrow \infty$  that

$$\begin{aligned} n^{-1/2} \sum_{i=1}^n \varphi_{f_0}(\Theta_i - \hat{\mu}^{(n)}) &= \Delta_{f_0}^{(n)}(\mu) - \mathbf{E}_{g_0}[\dot{\varphi}_{f_0}(\Theta_i - \mu)] n^{1/2}(\hat{\mu}^{(n)} - \mu) + o_{\mathbf{P}}(1) \\ &= \Delta_{f_0}^{(n)}(\mu) - \Gamma_{f_0,g_0;11} n^{1/2}(\hat{\mu}^{(n)} - \mu) + o_{\mathbf{P}}(1). \end{aligned}$$

(12)

Therefore, using (4) with (12), it follows that

$$\begin{aligned} \Delta_{f_0,j}^{(n)\star}(\hat{\mu}^{(n)}) &= \Delta_j^{(n)}(\mu) - \eta \Delta_{f_0}^{(n)}(\mu) - (\Gamma_{g_0,j;12} - \eta \Gamma_{f_0,g_0;11}) n^{1/2}(\hat{\mu}^{(n)} - \mu) + o_{\mathbf{P}}(1) \\ &= \Delta_{f_0,j}^{(n)\star}(\mu) + o_{\mathbf{P}}(1) \end{aligned}$$

under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$  as  $n \rightarrow \infty$  since  $\Gamma_{g_0,j;12} - \eta \Gamma_{f_0,g_0;11} = 0$ . It remains to show that

$$\hat{\Delta}_{f_0,j}^{(n)\star}(\hat{\mu}^{(n)}) - \Delta_{f_0,j}^{(n)\star}(\hat{\mu}^{(n)}) = - \left( \frac{\hat{\Gamma}_{g_0,j;12}}{\hat{\Gamma}_{f_0,g_0;11}} - \frac{\Gamma_{g_0,j;12}}{\Gamma_{f_0,g_0;11}} \right) \Delta_{f_0}^{(n)}(\hat{\mu}^{(n)}) = o_{\mathbb{P}}(1)$$

(13)

under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$  as  $n \rightarrow \infty$ . To prove (13), first note that (12) and the central limit theorem (CLT) imply that  $\Delta_{f_0}^{(n)}(\hat{\mu}^{(n)})$  is  $O_{\mathbb{P}}(1)$  under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$  as  $n \rightarrow \infty$ . Therefore, we only need to show that

$$\frac{\hat{\Gamma}_{g_0,j;12}}{\hat{\Gamma}_{f_0,g_0;11}} - \frac{\Gamma_{g_0,j;12}}{\Gamma_{f_0,g_0;11}} = o_{\mathbb{P}}(1)$$

(14)

as  $n \rightarrow \infty$  under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$  as  $n \rightarrow \infty$ . Since

$$\frac{\hat{\Gamma}_{g_0,j;12}}{\hat{\Gamma}_{f_0,g_0;11}} - \frac{\Gamma_{g_0,j;12}}{\Gamma_{f_0,g_0;11}} = \frac{\hat{\Gamma}_{g_0,j;12} - \Gamma_{g_0,j;12}}{\hat{\Gamma}_{f_0,g_0;11}} - \frac{\Gamma_{g_0,j;12} (\hat{\Gamma}_{f_0,g_0;11} - \Gamma_{f_0,g_0;11})}{\hat{\Gamma}_{f_0,g_0;11} \Gamma_{f_0,g_0;11}},$$

the result follows directly from Lemma 1.

Turning to the proof of (ii), and working along the same lines as those at the end of the proof of Lemma 1, it is easily shown that

$$n^{-1} \sum_{i=1}^n \varphi_{f_0}^2(\Theta_i - \hat{\mu}^{(n)}) - \mathbb{E}_{g_0}[\varphi_{f_0}^2(\Theta_i - \mu)],$$

$$n^{-1} \sum_{i=1}^n \sin^2(j(\Theta_i - \hat{\mu}^{(n)})) - \mathbb{E}_{g_0}[\sin^2(j(\Theta_i - \mu))]$$

(15)

and



$$n^{-1} \sum_{i=1}^n \sin(j(\Theta_i - \hat{\mu}^{(n)})) \varphi_{f_0}(\Theta_i - \hat{\mu}^{(n)}) - \mathbb{E}_{g_0} [\sin(j(\Theta_i - \mu)) \varphi_{f_0}(\Theta_i - \mu)]$$

(16)

are  $o_{\mathbb{P}}(1)$  under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$  as  $n \rightarrow \infty$ . Using the law of large numbers,  $C_{f_0,j}(\mu)$  converges to  $\mathbb{E}_{g_0} \left[ \left( \sin(j(\Theta_i - \mu)) - \frac{\Gamma_{g_0,j;12}}{\Gamma_{f_0,g_0;11}} \varphi_{f_0}(\Theta_i - \mu) \right)^2 \right]$ , in probability, under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$ . From this last result, it follows that  $C_{f_0,j}(\hat{\mu}^{(n)}) - C_{f_0,j}(\mu) = o_{\mathbb{P}}(1)$  under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$  as  $n \rightarrow \infty$ . Therefore it remains to show that  $\hat{C}_{f_0,j}(\hat{\mu}^{(n)}) - C_{f_0,j}(\hat{\mu}^{(n)})$  is  $o_{\mathbb{P}}(1)$  under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$  as  $n \rightarrow \infty$ . We readily obtain that

$$\begin{aligned} \hat{C}_{f_0,j}(\hat{\mu}^{(n)}) - C_{f_0,j}(\hat{\mu}^{(n)}) &= \left( \frac{\hat{\Gamma}_{g_0,j;12}^2}{\hat{\Gamma}_{f_0,g_0;11}^2} - \frac{\Gamma_{g_0,j;12}^2}{\Gamma_{f_0,g_0;11}^2} \right) n^{-1} \sum_{i=1}^n \varphi_{f_0}^2(\Theta_i - \hat{\mu}^{(n)}) \\ &\quad - 2 \left( \frac{\hat{\Gamma}_{g_0,j;12}}{\hat{\Gamma}_{f_0,g_0;11}} - \frac{\Gamma_{g_0,j;12}}{\Gamma_{f_0,g_0;11}} \right) n^{-1} \sum_{i=1}^n \sin(j(\Theta_i - \hat{\mu}^{(n)})) \varphi_{f_0}(\Theta_i - \hat{\mu}^{(n)}) \end{aligned}$$

so (15) and (16) together with (14) and the continuous mapping theorem imply that  $\hat{C}_{f_0,j}(\hat{\mu}^{(n)}) - C_{f_0,j}(\hat{\mu}^{(n)})$  is  $o_{\mathbb{P}}(1)$  under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$  as  $n \rightarrow \infty$ . The result follows.

□

## Proof of Theorem 2

Fix  $g_0 \in \mathcal{G}$  and  $\mu \in [-\pi, \pi)$ . Lemma 2 combined with Slutsky's lemma leads to

$$Q_{f_0,j}^{(n)} = \frac{\hat{\Delta}_{f_0,j}^{(n)*}(\hat{\mu}^{(n)})}{\sqrt{\hat{C}_{f_0,j}(\hat{\mu}^{(n)})}} = \frac{\Delta_{f_0,j}^{(n)*}(\mu)}{\sqrt{C_{f_0,j}(\mu)}} + o_{\mathbb{P}}(1) = \frac{\Delta_{f_0,j}^{(n)*}(\mu)}{V_{f_0}^{g_0}(j)} + o_{\mathbb{P}}(1)$$

(17)

as  $n \rightarrow \infty$  under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$ . Part (i) then follows from the CLT.

Part (ii) is obtained via Le Cam's third lemma. First, it is necessary to calculate the joint distribution of  $\hat{\Delta}_{f_0,j}^{(n)\star}(\hat{\mu}^{(n)})$  and  $\log(d\mathbf{P}_{(\mu,n^{-1/2}\tau_2^{(n)})';g_0,k}^{(n)}/d\mathbf{P}_{(\mu,0)';g_0}^{(n)})$  under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$ . We use Lemma 2 and the fact that

$$n^{-1/2} \sum_{i=1}^n \begin{pmatrix} \sin(j(\Theta_i - \mu)) - \eta\varphi_{f_0}(\Theta_i - \mu) \\ \tau_2^{(n)} \sin(k(\Theta_i - \mu)) \end{pmatrix} - \begin{pmatrix} 0 \\ \frac{1}{2}(\tau_2^{(n)})^2 \Gamma_{g_0,k;22} \end{pmatrix} \xrightarrow{\mathcal{D}} \mathcal{N}_2 \left( \begin{pmatrix} 0 \\ -\frac{1}{2}(\tau_2)^2 \Gamma_{g_0,k;22} \end{pmatrix}, \begin{pmatrix} V_{f_0}^{g_0}(j) & \tau_2 C_{f_0}^{g_0}(j,k) \\ \tau_2 C_{f_0}^{g_0}(j,k) & (\tau_2)^2 \Gamma_{g_0,k;22} \end{pmatrix} \right)$$

as  $n \rightarrow \infty$  under  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$ , obtained using the multivariate CLT. Now, since  $\mathbf{P}_{(\mu,0)';g_0}^{(n)}$  and  $\mathbf{P}_{(\mu,n^{-1/2}\tau_2^{(n)})';g_0,k}^{(n)}$  are mutually contiguous, applying Le Cam's third lemma, we obtain that  $\hat{\Delta}_{f_0,j}^{(n)\star}(\hat{\mu}^{(n)}) \xrightarrow{\mathcal{D}} \mathcal{N}(\tau_2 C_{f_0}^{g_0}(j,k), V_{f_0}^{g_0}(j))$  under  $\mathbf{P}_{(\mu,n^{-1/2}\tau_2^{(n)})';g_0,k}^{(n)}$ , as  $n \rightarrow \infty$ .

We now turn to Part (iii). From (17), it is easily seen that

$$n^{-1/2} \sum_{i=1}^n \left[ \sin(j(\Theta_i - \mu)) - \frac{\Gamma_{f_0,j;12}}{\Gamma_{f_0,f_0;11}} \varphi_{f_0}(\Theta_i - \mu) \right] = \Delta_{f_0,j}^{(n)\star}(\hat{\mu}^{(n)}) + o_{\mathbf{P}}(1), \quad (18)$$

as  $n \rightarrow \infty$  under  $\mathbf{P}_{(\mu,0)';f_0}^{(n)}$ , and therefore under contiguous alternatives. Now, consider the following parametric testing problem

$$\begin{cases} \mathcal{H}_{0;f_0} := \cup_{\mu \in [-\pi, \pi]} \mathbf{P}_{(\mu,0)';f_0}^{(n)}, \\ \mathcal{H}_{1;f_0,j} := \cup_{\lambda \neq 0 \in [-1,1]} \cup_{\mu \in [-\pi, \pi]} \mathbf{P}_{(\mu,\lambda)';f_0,j}^{(n)}, \end{cases}$$

and consider the parametric test that rejects  $\mathcal{H}_{0;f_0}$  when the absolute value of the parametric test statistic, that can be derived by taking as numerator the first part of the equality in (18), is greater than  $z_{1-\alpha/2}$ . It follows from Le Cam theory that this test is locally and asymptotically maximin for testing  $\mathcal{H}_{0;f_0}$  against  $\mathcal{H}_{1;f_0,j}$ . The result in (iii) then follows from (18) and the optimality features of the parametric test for all  $f_0 \in \mathcal{G}$  satisfying Assumption A.  $\square$

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