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# Has the Pandemic Triggered a 'Paperdemic'? Towards an Assessment of Diagnostic Indicators for COVID-19

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### Authors' contributions

This work was carried out in collaboration between the two authors. Author AMAR wrote the entire draft of the manuscript, conducted the mathematical and conceptual analyses and managed the basic literature survey. Author HAMS participated in the literature search, performed the computational work and constructed the table of results. Both authors read and approved the final manuscript.

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### **ABSTRACT**

This paper is a preliminary step towards the assessment of an alarming widespread belief that victims of the novel coronavirus SARS-CoV-2 include the quality and accuracy of scientific publications about it. Our initial results suggest that this belief cannot be readily ignored, denied, dismissed or refuted, since some genuine supporting evidence can be forwarded for it. This evidence includes an obvious increase in retractions of papers published about the COVID-19 pandemic plus an extra-ordinary phenomenon of inconsistency that we report herein. In fact, we provide a novel method for validating any purported set of the four most prominent indicators of diagnostic testing (Sensitivity, Specificity, Positive Predictive Value, and Negative Predictive Value), by observing that these indicators constitute three rather than four independent quantities. This observation has virtually been unheard of in the

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open medical literature, and hence researchers have not taken it into consideration. We define two functions, which serve as consistency criteria, since each of them checks consistency for any set of four numerical values (naturally belonging to the interval [0.0,1.0]) claimed to be the four basic diagnostic indicators. Most of the data we came across in various international journals met our criteria for consistency, but in a few cases, there were obvious unexplained blunders. We explored the same consistency problem for some diagnostic data published in 2020 concerning the ongoing COVID-19 pandemic and observed that the afore-mentioned unexplained blunders tended to be on the rise. A systematic extensive statistical assessment of this presumed tendency is warranted.

Keywords: COVID-19; compromised standards; diagnostic testing; sensitivity; specificity; positive predictive value; negative predictive value; consistency criterion.

### 1. INTRODUCTION

The world currently witnesses an ongoing epidemic of the novel coronavirus (SARS-CoV-2) that causes the disease COVID-19, now characterized as a pandemic by the World Health Organization (WHO) [1-6]. This pandemic seems to have expanded from the Wuhan province in China, but has definitely reached (in repetitive waves) almost every inhabited territory on the globe. This fatal disease, being a catastrophic threat of dramatic public-health and economic concerns proved to have diverse grave (and potentially irreversible) consequences [1-4]. For the past thirty years, a once-per-decade novel coronavirus has pushed the global public health system to the limit, with SARS-CoV-2 being the most severe. Despite repeated warnings, the world was not prepared for that pandemic [7], and the world's response to the pandemic leaves a lot to be desired. Now, there is a genuine need for further research concerning various aspects of epidemiology, in general, with a stress on pathogen research related to COVID-19, in particular. This research might pay off handsomely in improving the response of humanity to this pandemic. The past year (2020) witnessed a dramatic increase in the number of published scientific papers supposedly reporting results of the required research. There is a great concern that publication standards have been compromised as the peer-review process is becoming hasty and weak [8-16]. There are suspicions that scientific publications concerning COVID-19 are ranging from robust and rigorous studies to dishonest, incompetent, or fraudulent studies being conducted, posted, and shared at an unprecedented rate. Dinis-Oliveira [8] coined the term 'paperdemic' to refer to a parallel virtual viral pandemic (ignited by the genuine pathogenic pandemic) that leads to publications based on mistakes or misconduct. He asserts that A pandemic with a "paperdemic" will be even more complicated to manage if it progresses in

an uncontrolled manner and is not properly scrutinized.

The role of gatekeepers of science is the collective responsibility of scientists in general [17]. It is also an implicit assignment or task for the contemporary system of scientific publishing. This system addresses such a task in a proactive strategy (the peer-review process) supplemented by a reactive manner (paper retraction schemes). Since we are suspicious of the proactive strategy in the COVID-19 era [18-20], we will look now at the reactive paradigm to see if it supports the hypothesized phenomenon of compromised standards and if it can serve to mitigate it. Paper retraction might only address severe issues such as fraud (author engagement in research misconduct or author use of deliberately flawed, fabricated, falsified or fraudulent data) or errors (plagiarism or scientific mistakes). Retractions are in some cases accompanied by explanations and apologies for what warrants the total content withdrawal (rather than a limited-scope correction). A decade ago, Steen [21] observed that the recent rate of increase in retractions was generally greater than the corresponding rate of increase in publications. Moreover, he noted that in the 6 years between 2004 and 2009. the increase in the number of papers retracted for fraud was much more than that retracted for scientific mistakes [21]. Vuong [22] asserts that retractions are not intrinsically bad, as they serve as a practical way to correct for human fallibility and to keep the scientific literature trustworthy. He points out that the scientific community should agree on the essential information to be provided when pulling a paper from the scientific literature. Definitely, retraction notices should be more informative [22,23]. During the COVID-19 pandemic, an exceptionally fast track became available for the peer review of papers dealing with the pandemic. The availability of such a fast track is blamed for further increase in the number of papers that have been retracted due to quality

and/or data issues [22,24-26]. In fact, such a fast track is widely accused of being of low-quality despite vigorous denial by publishers.

Even before the COVID-19 era, there have been many methods (mainly statistical) for detecting false data [27-37], that have become rather visible and advanced during the past decade. Rushdi and Rushdi [38,39] have recently suggested methods for avoiding probabilistic fallacies in medical context. In [40], they introduced a non-statistical method based on the premise that flawed data might be detected via the excessive inconsistencies it causes in a variant of Boolean Analysis called Qualitative Comparative Analysis (QCA) [40-44]. Rushdi and Serag [4,45] developed yet another nonstatistical checking method, which can be used for validating a certain category of bio-statistical data. This method is based on a newlydiscovered inter-relation among the four most prominent indicators of diagnostic testing (sensitivity, specificity, positive predictive value (PPV), and negative predictive value (NPV)) [4,40,45-49]. Rushdi and Serag [4,45] developed simple formulas that express any one of these four indicators in terms of the other three. They called a set of four values satisfying these formulas (to within permissible round-off errors) a consistent set. Extensive testing was made for sets of the four basic indicators published recently in a variety of international medical journals [45], and in various Saudi medical journals [17], to check whether these sets are consistent or not. Most of the data they came across met their criteria for consistency, but in a few cases, there were obvious unexplained blunders. The present paper aims to explore the same consistency problem for some diagnostic data published in 2020 concerning the ongoing COVID-19 pandemic. We observe that the aforementioned unexplained blunders tended to be on the rise within publications about COVID-19, and we suggest that a systematic extensive statistical assessment of this presumed tendency is warranted.

The organization of the rest of this paper is as follows. Section 2 is a standard section of materials and methods. It starts with a brief primer about diagnostic testing and its basic measures. Section 2 also presents the essential machinery on which this paper is based. It reports virtually unknown formulas for interdependence among the two predictive values, sensitivity, and specificity. These formulas express any one of these four indicators in terms

of the other three, under the assumption that each of the four exists, and no division by zero is encountered. Section 3 reports our computational results, obtained by applying the new formulas extensively to some data published in 2020 about diagnostic measures for COVID-19. Most sets of values of sensitivity, specificity, and predictive values tested agree with our formulas, thereby independently attesting to the correctness of these formulas. However, some reported sets of the four basic indicators experience some appreciable incoherence among their values according to our formulas. The percentage of incoherent cases seems definitely higher than it was in [17,45]. Section 4 presents a detailed discussion of the importance of our findings, while Section 5 concludes the paper.

### 2. MATERIALS AND METHODS

## 2.1 On Diagnostic Testing and Its Basic Measures

This subsection is intended for a brief primer about diagnostic testing and its most basic indicators [38-40,45-49]. Fig. 1 demonstrates a two-by-two contingency matrix for test or classification i with respect to test classification j. Each of the two variables i and j is a dichotomous variable that belongs to the set  $\{+1,-1\}$  of indices. The test *i* reports 'positive' cases (arbitrarily assigned the value +1), in which a certain disease, attribute, trait, or condition is present, or reports 'negative' cases (arbitrarily assigned the value -1), in which this disease, attribute, trait, or condition is absent. This test is assessed or evaluated by a reference or gold standard test j, which has its own labeling of cases, again as positive or negative. The reference test *i* designates various cases of the assessed test i as "true" or "false." depending on whether it agrees or disagrees with test i, respectively. As a result, the matrix four entries are called True Positives, False Positives, False Negatives, and True Negatives. These entries are usually assigned the standard abbreviations TP, FP, FN, and TN. In the sequel, we will use the subscripted abbreviations  $TP_{ij}$ ,  $FP_{ij}$ ,  $FN_{ij}$ , and  $TN_{ij}$ , where we use the subscripts ij for all measures (and later for indicators derived from them) to assert the notion that i is assessed, judged or measured relative to i. The sum of these four entries is the size of the reported population or the total number of reported cases N. If the tests i and j interchange their roles (so that test j is now assessed lative to test i) then the four measures are relabeled as  $TP_{ji}$ ,  $FP_{ji}$ ,  $FN_{ji}$ , and  $TN_{ji}$  such that  $TP_{ji} = TP_{ij}$ , and  $TN_{ji} = TN_{ij}$  but with  $FP_{ji} = FN_{ij}$ , and  $FN_{ji} = FP_{ij}$ . This is the reason why omission of the subscripts is not desirable, as it leads to an inadvertent ambiguity as to which assesses which.

We use the symbols  $A = \{j = +1\}$  and  $B = \{i = 1\}$ +1} to denote the events of positive cases (presence of the considered condition) according to the tests j and i, respectively. Hence, the complementary events  $\overline{A} = \{j = -1\}$  and  $\overline{B} = \{i = 1\}$ -1} denote the events of negative cases (absence of the considered condition) according to the tests j and i, respectively. There are eight conditional probabilities concerning these two events and their complements, as shown in Fig. 2. These can be identified as the eight most prominent indicators used in diagnostic testing. These are the Sensitivity (Sensii) or True Positive Rate  $(TPR_{ii})$ , the Specificity  $(Spec_{ii})$  or True Negative Rate (  $TNR_{ij}$  ), the Positive and Negative Predictive Values ( $PPV_{ij}$  and  $NPV_{ij}$ ), together with their respective complements (to 1.0), namely the False Negative Rate (FNRii), False Positive Rate ( $FPR_{ij}$ ), False Discovery rate  $(FDR_{ii})$  and False Omission Rate  $(FOR_{ii})$  [38-40,45-49]. The former four indicators are considered more popular or more prominent, and they act as direct or agreement measures the latter four serve as discrepancy or disagreement measures between the two tests i and j. Due to the four complementation relations within pairs of these eight measures, the number of independent quantities among them is at most four. It seems that there is a widespread (and at least implicit) belief that this number is exactly four (usually obtained by counting the four direct indicators  $Sens_{ii}, Spec_{ii}, PPV_{ii}$  and  $NPV_{ii}$ ). We show in Section 3 that this number is, in fact, three, by simply being able to express any of the four direct indicators in terms of the other three.

Note that each conditional probability in Fig. 2 has a 'dual' one obtained by complementing both the conditioned and conditioning events [50], and also has an inverse or transposed one, obtained by swapping or interchanging the conditioned and conditioning events [38,39]. Our definition of 'duality' is in line with that used with Boolean quantities, where duality is achieved through complementing both the input and output quantities [51-53]. Our definitions of duality and transposition mean that each conditional

probability P has a dual  $P^d$ , a transpose or inverse T, and a dual of its transpose or inverse (a transpose of its dual)  $T^d$ . Note that both the duality and transposition operators are involutary or self-inverse operators, i.e., each of them satisfies 'the law of involution' (applying any of them twice to a specific conditional probability leaves it intact) [54-56]. Table 1 defines the four possible sets  $\{P, P^d, T, T^d\}$  pertaining to the set of four direct indicators of diagnostic testing. Similar definitions apply to the set of complementary indicators of diagnostic testing. Fig. 3 is yet another geometric display of the inter-relationships among the eight diagnostic indicators defined in Fig. 2. Two conditional probabilities constituting a dual pair are placed on the same vertical line, while two conditional probabilities constituting a transpose or inverse pair are situated on the same horizontal line. Hence, any conditional probability and the dual of its transpose or inverse (the transpose of its dual) appear diagrammatically opposite. In Fig. 3, the set of four direct indicators of diagnostic testing is distinguished in blue, while the set of complementary indicators of diagnostic testing is highlighted in red. Each member in the first set has a one-to-one and onto mapping to a member in the second set, which is its complement (to one).

### 2.2 Validating Formulas Used in the Analysis

We now express each of the four most prominent indicators of diagnostic testing (Specificity, Negative Predictive Value, Sensitivity, and Positive Predictive Value) solely in terms of the other three (provided each of the four indicators exists, and no division by zero is encountered), namely [4,17,45].

$$Sens_{ij} = \frac{PPV_{ij} * NPV_{ij} [1 - Spec_{ij}]}{PPV_{ij} NPV_{ij} + Spec_{ij} [1 - PPV_{ij} - NPV_{ij}]}$$
(1)

$$Spec_{ij} = \frac{PPV_{ij} * NPV_{ij} \left[1 - Sens_{ij}\right]}{PPV_{ij} * NPV_{ij} + Sens_{ij} \left[1 - PPV_{ij} - NPV_{ij}\right]}$$
(2)

$$PPV_{ij} = \frac{Sens_{ij} * Spec_{ij} \left[1 - NPV_{ij}\right]}{Sens_{ij} * Spec_{ij} + NPV_{ij} \left[1 - Sens_{ij} - Spec_{ij}\right]}$$
(3)

$$NPV_{ij} = \frac{Sens_{ij} * Spec_{ij} \left[1 - PPV_{ij}\right]}{Sens_{ij} * Spec_{ij} + PPV_{ij}\left[1 - Sens_{ij} - Spec_{ij}\right]}$$
(4)

Equations (1-4) might be written in a unified form (See Table 1) as

$$P = \frac{T * T^{d} [1 - P^{d}]}{T * T^{d} + P^{d}[1 - T - T^{d}]}$$
 (5)

We also define two checking functions of these four values that we call the Diagnostic Checking Difference (DCD) and the Diagnostic Checking Ratio (DCR), that are exactly 0 and 1, respectively, for consistent values. The mathematical definition of the DCD and DCR is [4].

$$DCD_{ij} = Sens_{ij} * Spec_{ij} [PPV_{ij} + NPV_{ij} - 1-PPV_{ij} * NPV_{ij} [Sens_{ij} + Spec_{ij} - 1],$$
 (6)

$$DCR_{ij} = \frac{Sens_{ij} * Spec_{ij} \left[ PPV_{ij} + NPV_{ij} - 1 \right]}{PPV_{ij} * * NPV_{ij} \left[ Sens_{ij} + Spec_{ij} - 1 \right]}.$$
 (7)

We reiterate that we use the subscripts ij for all measures and indicators to assert the notion that test i is assessed, judged or measured relative to the reference test or gold standard j. Equations (6) and (7) might be written in a generalized form (See Table 1) as.

$$+DCD_{ij} OR -DCD_{ij} = P * P^{d} [T + T^{d} - 1] - T * T^{d} [P + P^{d} - 1].$$
 (8)

$$DCR_{ij} OR (1/DCR_{ij}) = \frac{P * P^{d} [T + T^{d} - 1]}{T * T^{d} [P + P^{d} - 1]}.$$
 (9)

The fact that the expressions in (8) and (9) are identically equal to 0 and 1, respectively, means that the quantity  $(P*P^d[T+T^d-1])$ , which is naturally invariant to the replacement of every term by its dual, is also invariant to the replacement of every term by its transpose.

### 3. RESULTS

This section reports our results, followed by an assessment of some (arbitrary selected) diagnostic data reported in the COVID-19 Era. We note that the deviation of the DCD and the DCR from 0 and 1, respectively, is a measure of inconsistency for any purported set of the four

diagnostic indicators (sensitivity, specificity, positive predictive value (PPV), and negative predictive value (NPV)). Since members of such a purported set are (conditional) probabilities, they are expected to be non-negative fractional vales belonging to the interval [0.0, 1.0]. Table 2 provides our validation of some published sets of these four basic indicators within some publications that appeared in 2020 on the COVID-19 pandemic. We check whether the sets considered are consistent or not. For each published set of  $\{Sens_{ij}, Spec_{ij}, PPV_{ij}, NPV_{ij}\}$  the table computes the checking difference  $DCD_{ij}$  via (6), and the checking ratio DCRii via (7). It also uses equations (1-4) to compute a new value for each of the four prominent indicators in terms of the old values of the other three indicators. We arbitrarily assume that a published set is consistent (uncolored entries) if the absolute value of the relative error is less than or equal to 2%. We arbitrarily consider such a small error accountable for by normal or acceptable round-off errors [57-62]. Otherwise, we consider a set to be somewhat problematic (with error still within 4%, highlighted in yellow), or inconsistent (with error still within 6%, highlighted in orange). If the absolute relative error exceeds 6%, we arbitrarily label the corresponding set as dramatically inconsistent (highlighted in Most sets of values of sensitivity, red). specificity, and predictive values tested agree with our formulas, thereby independently attesting to the correctness of these formulas. However, some reported sets of the four basic some indicators experience appreciable incoherence among their values according to our formulas. Similar results were earlier obtained in [45] for data randomly selected from international journals and in [17] for data randomly selected from Saudi medical journals. The percentage of incoherent cases and the severity of their inconsistency seem definitely higher for the present COVID-19 data than they were in [17,45]. In fact, less than half of the articles covered in Table 2 are totally free of any inconsistency problem.

### 4. DISCUSSION

Our results could become a useful future methodology to verify a certain aspect of the quality of medical articles. This methodology would complement other existing guidelines for the reporting of medical research [63-71]. The purpose of having these guidelines is to establish a recipe or "manual for the authors to follow, which should lead to total transparency, accurate

reporting, and easier assessment of the validity of reported research findings [69]." The six most widely accepted and used guidelines are: CONSORT, STROBE, PRISMA. MOOSE. STARD, and SPIRIT. Johansen and Thomsen [69] assert that the implementation of these guidelines has led to only a moderate improvement in the quality of the reporting of medical research, and that there is still much work to be done to achieve accurate and transparent reporting of medical research findings. We hope the validating formulas reported herein would receive the popularization they deserve, and that future medical publications would take them into consideration. We do not anticipate a dramatic improvement in

the quality of reporting of medical research, but we hope that our humble contribution would serve as a modest step of a process of continuous improvement. Moher and Altman [67] assert that "making a major impact on the quality of reporting and mitigating deficiencies is a huge challenge because no one group has prime responsibility and no single action is likely to have a large impact." We agree with the four proposals they made to help improve the medical research literature, namely: "(1) introducing publications officers; (2) developing core competencies for editors and peer reviewers, around which (3) training can be tailored; and (4) training authors to write articles fit for purpose."

i	+1	-1
+1	<i>TP<sub>ij</sub></i> (True Positives)	FP <sub>ij</sub> (False Positives) (Type I Error)
-1	FN <sub>ij</sub> (False Negatives) (Type II Error)	<i>TN<sub>ij</sub></i> (True Negatives)

Fig. 1. The two-by-two contingency matrix of test or classification i with respect to test or classification j. This matrix has integer entries that add to the total number of cases N. The symbols  $A = \{j = +1\}$  and  $B = \{i = +1\}$  denote the events of positive cases according to tests j and i, respectively

Table 1. Possible definitions of a conditional probability P, its dual  $P^d$ , its transpose or inverse T, and the dual of its transpose or inverse (transpose of its dual)  $T^d$ . These definitions pertain to the set of four direct indicators of diagnostic testing. Similar definitions apply to the set of complementary indicators of diagnostic testing

P	P <sup>d</sup>	T	T <sup>d</sup>	
Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	$NPV_{ij}$	
Spec <sub>ij</sub>	Sens <sub>ij</sub>	$NPV_{ij}$	$PPV_{ij}$	
$PPV_{ij}$	$NPV_{ij}$	Sens <sub>ij</sub>	$Spec_{ij}$	
$NPV_{ij}$	$PPV_{ij}$	Spec <sub>ij</sub>	Sens <sub>ij</sub>	

			B conditioned	
	$P(\bar{A} \bar{B}) = P(j = -1 i = -1)$ $= NPV_{ij}$	$P(A \bar{B}) = P(j = +1 i = -1)$ $= FOR_{ij}$	$P(B \bar{A}) = P(i = +1 j = -1)$ $= FPR_{ij}$	$P(\bar{B} \bar{A}) = P(i = -1 j = -1)$ $= Spec_{ij} = TNR_{ij}$
Conditioning uncomplemented	$P(\bar{A} B) = P(j = -1 i = +1)$ $= FDR_{ij}$	$P(A B) = P(j = +1 i = +1)$ $= PPV_{ij}$	$P(B A) = P(i = +1 j = +1)$ $= Sens_{ij} = TPR_{ij}$	$P(\bar{B} A) = P(i = -1 j = +1)$ $= FNR_{ij}$
		Conditioned uncomple	mented	

Fig. 2. Definition of the eight conditional probabilities concerning events  $A = \{j = +1\}$  and  $B = \{i = +1\}$ , which constitute the eight most prominent indicators of diagnostic testing. The four shaded entries are direct indicators, usually taken for the most basic ones. The four unshaded entries are complementary indicators. Each conditional probability has a 'dual' one obtained by complementing both the conditioned and conditioning events, and also has an inverse or transposed one, obtained by swapping the conditioned and conditioning events

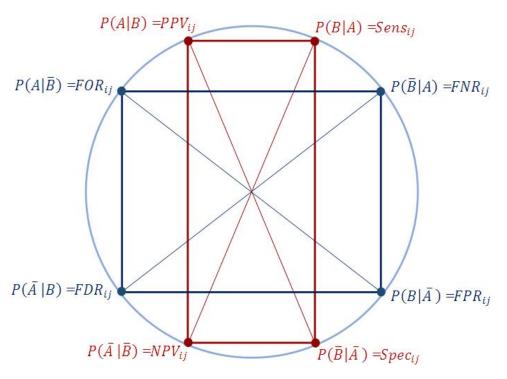


Fig. 3. Geometric display of the inter-relationships among the eight diagnostic indicators defined in Fig. 2. Each of these indicators is a conditional probability P that has a dual  $P^d$  (on the same vertical line), a transpose or inverse T (on the same horizontal line), and a dual of its transpose or inverse (a transpose of its dual)  $T^d$  (diagrammatically opposite). The set of four direct indicators of diagnostic testing is distinguished in blue, while the set of complementary indicators of diagnostic testing is highlighted in red

Table 2. Checking consistency among sets of the four prominent diagnostic indicators published in 2020 on topics related to the Covid-19 pandemic. In a dominant majority of cases, the published sets are consistent (uncolored entries), and in a small number of cases, there are sets that are somewhat problematic (highlighted in yellow), or dramatically inconsistent (highlighted in red)

#		Origina	l values	i		cking lues		Compute	ed values	3	Source	Data location
	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	$NPV_{ij}$	DCD <sub>ij</sub>	DCR <sub>ij</sub>	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	NPV <sub>ij</sub>		in Source
1					0.0001		0.8903	0.9909	0.8156	0.9934	Manski, C. F. (2020). Bounding the Predictive Values of COVID-	
				0.9970		1.0003	0.9446			0.9966	19 Antibody Tests (No. w27226). National Bureau of Economic	
	0.8800	0.9880	0.7940	0.9940	0.0001	1.0001	0.8858	0.9886	0.7847	0.9937	Research. Available at:	
											https://www.nber.org/system/files/working_papers/w27226/w2722 6.pdf.	
2					0.0000			0.7804	0.5064	0.9850	Ji D, Zhang D, Xu J, Chen Z, Yang T, Zhao P, Chen G, Cheng G,	Table 4
	0.4500	0.9700	0.7830	0.1190	-0.0819	-1.0931	0.0148	0.3733	0.9949	0.8800	Wang Y, Bi J, Tan L. Prediction for progression risk in patients	
											with COVID-19 pneumonia: the CALL score. Clinical Infectious	
	0.7700	4.0000	4.0000	0.0000	0.0000	4.0000	#DI) ((0)	4.0000	4.0000	#DI) (/OI	Diseases. 2020;71(6):1393-1399.	
3				0.8000		1.0000 0.9999	#DIV/0! 0.8323	1.0000 0.9498	1.0000 0.9482	#DIV/0! 0.8387	Xiang F, Wang X, He X, Peng Z, Yang B, Zhang J, Zhou Q, Ye H, Ma Y, Li H, Wei X. Antibody detection and dynamic characteristics	
				0.0300			#DIV/0!	1.0000	1.0000	#DIV/0!	in patients with coronavirus disease 2019. Clinical Infectious	
					-0.0037		0.6198	0.9503	0.8939	0.9240	Diseases. 2020;71(8):1930-1934.	
4	0.9176	1.0000		0.9739			#DIV/0!	1.0000	1.0000	#DIV/0!	Bisoffi Z, Pomari E, Deiana M, Piubelli C, Ronzoni N, Beltrame A,	Fig. 3
	0.7619	0.9962	0.9846	0.9286	0.0000	1.0000	0.7603	0.9962	0.9847	0.9292	Bertoli G, Riccardi N, Perandin F, Formenti F, Gobbi F.	3 -
	0.6118	0.9962	0.9811	0.8874	0.0000	1.0000	0.6095	0.9962	0.9813	0.8884	Sensitivity, specificity and predictive values of molecular and	
				0.9811		1.0000	0.9415	0.9923	0.9755	0.9810	serological tests for COVID-19: a longitudinal study in emergency	
				0.9804		1.0000	0.9411	0.9578	0.8792	0.9804	room. Diagnostics. 2020;10(669):1-12.	
				0.9736		1.0000	0.9168	0.9885	0.9625	0.9736		
	0.8941			0.9650		1.0000	0.8941	0.9502	0.8539	0.9650		
				0.9350		1.0000	0.7866 0.6243	0.9923 0.9923	0.9705 0.9635	0.9347 0.8896		
				0.8300		1.0000	0.0243	0.9923	0.9033	0.8389		Fig. 5
				0.8225			0.4237	0.8698	0.5141	0.8224		1 ig. 5
				0.8197		0.9999	0.3763	0.9234	0.6156	0.8198		
	0.3765	0.9234	0.6154	0.8197	0.0000	0.9999	0.3763	0.9234	0.6156	0.8198		
				0.8022		1.0003	0.3530	0.8545	0.4411	0.8021		
				0.8176			0.3411	0.9617	0.7437	0.8177		
				0.8084		1.0000	0.3059	0.9540	0.6842	0.8084		
	0.2941			0.8083		1.0001	0.2945	0.9694	0.7573	0.8080		
	0.2941	0.9655	0.7353	0.8077	0.0000	1.0001	0.2942	0.9655	0.7352	0.8076		

#		Origina	l values			cking ues		Compute	ed values	<b>3</b>	Source	Data location
	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	$NPV_{ij}$	$DCD_{ij}$	DCR <sub>ij</sub>	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	$NPV_{ij}$		in Source
	0.2824	0.9464	0.6316	0.8019	0.0000	0.9998	0.2822	0.9463	0.6319	0.8021		
					0.0000	0.9999	0.2467	0.9847	0.8403	0.8009		
				0.7913		0.9999	0.2117		0.7202	0.7914		
				0.8611		1.0000	0.5714	0.8857	0.6250	0.8611		Fig. 6
				0.8545		1.0000	0.5428		0.6334	0.8545		
				0.8462		1.0000	0.5430	0.8381	0.5277	0.8462		
				0.8509		1.0000	0.5144	0.9238	0.6923	0.8509		
				0.8247		0.9999	0.5142	0.7618	0.4187	0.8247		
				0.8462		1.0000	0.4856	0.9429	0.7392	0.8463		
				0.8333		1.0001	0.4858	0.8571	0.5312	0.8332		
				0.8390		1.0000	0.4570	0.9429	0.7274	0.8390		
				0.8362		1.0000	0.4572		0.6666	0.8361		
				0.8211		1.0000	0.3714		0.7647	0.8211		
				0.8167		1.0001	0.3716		0.6498	0.8166		
				0.8095		1.0000	0.3145		0.7856	0.8094	0:00 01 01:71:07	
5				0.9400		1.0001	0.8502	0.7702	0.5477	0.9399	Cui S, Chen S, Li X, Liu S, Wang F. Prevalence of venous	Table 3
				0.9470		1.0010	0.8552		0.6993	0.9449	thromboembolism in patients with severe novel coronavirus	
				0.9320	-0.0001	1.0003	0.7986 0.7023	0.9012 0.9347		0.9325 0.9041	pneumonia. Journal of Thrombosis and Haemostasis.	
				0.9080		1.0003	0.7023	0.9347		0.9041	2020;18(6):1421-1424.	
				0.8940		1.0002	0.7022	0.9673		0.8930		
6				1.0000		1.0002	1.0000		#DIV/0!		Ni Q, Sun ZY, Qi L, Chen W, Yang Y, Wang L, Zhang X, Yang L,	Table 2
U				0.6200		1.0000	#DIV/0!	1.0000	1.0000	#DIV/0!	Fang Y, Xing Z, Zhou Z. A deep learning approach to characterize	Table 2
				0.5000		1.0048	0.9423		0.9755	0.4486	2019 coronavirus disease (COVID-19) pneumonia in chest CT	
				0.4400			#DIV/0!	1.0000	1.0000	#DIV/0!	images. European Radiology. 2020;30(12):6517-6527.	
					0.0014		0.9651		0.7547	0.9202	images. European Radiology. 2020,00(12).0017-0027.	
					-0.0007		0.8619	0.8935	0.9250	0.8397		
					-0.0010		0.8251		0.9458	0.8166		
						1.0009	0.8352	0.9639	0.9667	0.7718		
						1.0015	0.9626		0.7781	0.9253		Table 3
						1.0007	0.9154	0.9060	0.9254	0.8726		
					-0.0010		0.9139	0.9139	0.9515	0.9185		
				0.8900		1.0001	0.9109	0.9505	0.9596	0.8889		
				0.9100		1.0002	0.9404		0.7588	0.9095		
	0.8200	0.8900	0.8900	0.8200	0.0000	1.0000	0.8200	0.8900	0.8900	0.8200		

:	(	Origina	l values			cking ues		Compute		3	Source	Data location
Sen	S <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	$NPV_{ij}$	$DCD_{ij}$	DCR <sub>ij</sub>	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	$NPV_{ij}$		in Source
0.780	00	0.9600	0.9500	0.8000	-0.0008	0.9986	0.7600	0.9554	0.9551	0.8175		
				0.7800		1.0026	0.7800	0.9370		0.7600		
				0.9500		1.0046	0.9837	_	0.6872			
				0.8100			0.8656		0.9060	0.8026		
					-0.0003		0.8364	0.8976	0.9220	0.8043		
					0.0010	1.0018		0.9763	0.9746	0.7128		
					0.0010	1.0019		0.7231	0.7814			
					-0.0007		0.8829	0.8723				
					-0.0008		0.7600	0.9053	0.9145	0.7700	•	
					-0.0017		0.7128	0.9700	0.9800	0.8508		
					-0.0006			0.5937	0.8489	0.9060		
						1.0006		0.9434	0.9681	0.7485		
				0.7600		1.0006	0.8673	0.9434	0.9681	0.7485		
					-0.0008		0.8197	0.9629	0.9839	0.7890		<b></b> .
				0.6200		1.0000	#DIV/0!	1.0000	1.0000	#DIV/0!		Table 5
				0.8000		1.0000	#DIV/0!	1.0000	1.0000	#DIV/0!		
				0.5000		1.0048	0.9423		0.9755	0.4486		
					-0.0001		0.9692	0.8772	0.9903	0.7055		
				0.4400		1.0000	#DIV/0!	1.0000	1.0000	#DIV/0!		
				0.7300		1.0000	#DIV/0!	1.0000	1.0000	#DIV/0!		
					-0.0007		0.8619	0.8935	0.9250	0.8397		
					-0.0009		0.8963	0.9191	0.9482	0.8956		
					-0.0010		0.8251	0.9227	0.9458	0.8166		
					-0.0009		0.9246	0.9246	0.9604	0.9282		
				0.7900		1.0009 1.0004	0.8352	0.9639	0.9667	0.7718		
				0.8600	-0.0102		0.8858	0.9526	0.9578	0.8531 0.5657	Won 7 Chi V 7hang L Liu H Du K Li 7 8 Wang D (2020)	Table 3
					-0.0102	_				<u></u> .	Wen Z, Chi Y, Zhang L, Liu H, Du K, Li Z & Wang D. (2020).	Table 3
									0.9390	0.5958	Coronavirus disease 2019: initial detection on chest CT in a	
0.950	00	0.4300	0.9000	0.0000	-0.0009	0.9954	0.9471	0.4 154	0.9053	0.6143	retrospective multicenter study of 103 Chinese subjects.	
0.07	20	0.0500	0.6500	0.0200	0.0000	0.0007	0.0645	0.0400	0.0000	0.0520	Radiology: Cardiothoracic Imaging, 2(2), e200092.	Table 2
							0.9645		0.6882	0.8530	Ai T, Yang Z, Hou H, Zhan C, Chen C, Lv W & Xia L. (2020).	Table 2
					-0.0006		0.9586		0.6284	0.8447	Correlation of chest CT and RT-PCR testing in coronavirus	
					0.0015		0.9733	0.2413		0.7801	disease 2019 (COVID-19) in China: a report of 1014 cases.	
					-0.0002		0.9596		0.6524	0.7519	Radiology, 200642.	
0.970	JU	0.3000	0.0000	0.8900	0.0015	1.0092	0.9/34	0.3269	0.6314	0.8771		

#		Origina	l values			cking ues		Compute	d values	•	Source	Data location
	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	NPV <sub>ij</sub>	DCD <sub>ij</sub>	DCR <sub>ij</sub>	Sens <sub>ij</sub>	Spec <sub>ij</sub>	PPV <sub>ij</sub>	NPV <sub>ij</sub>		in Source
9					-0.0010		0.7822	0.8337	0.8848	0.7292	Guillo E, Gomez IB, Dangeard S, Bennani S, Saab I, Tordjman M	
					-0.0003		0.8062	0.9080	0.9316	0.7644	& Revel MP. (2020). COVID-19 pneumonia: Diagnostic and	
					0.0002		0.9013	0.7824		0.7977	prognostic role of CT based on a retrospective analysis of 214	
	0.9300	0.8800	0.9400	0.8600	-0.0001	0.9999	0.9292	0.8787	0.9407	0.8615	consecutive patients from Paris, France. European journal of radiology, 131, 109209.	
10	0.9000	0.9000	0.3210	0.9940	-0.0001	0.9996	0.8969	0.8969	0.3284	0.9942	Kumleben N, Bhopal R, Czypionka T, Gruer L, Kock R, Stebbing	Table 1
	0.8000	0.9900	0.8080	0.9890	-0.0001	0.9999	0.7926	0.9895	0.8150	0.9895	J, Stigler FL. Test, test, test for COVID-19 antibodies: the	Table 2
	0.9900	0.9900	0.8380	0.9990	-0.0001	0.9999	0.9812	0.9812	0.9075	0.9995	importance of sensitivity, specificity and predictive powers. Public Health. 2020;185:88-90.	Table 3
11	0.7330	0.4130	0.3330	0.7950	0.0001	1.0025	0.7335	0.4136	0.3325	0.7946	La Torre G, Massetti AP, Antonelli G, Fimiani C, Fantini M, Marte	Table 3
					-0.0001		0.5997	0.4667	0.3103	0.7452	M & Villari, P. (2020). Anosmia and ageusia as predictive signs of	
	0.1000	0.9870	0.7500	0.7327	-0.0002	0.9965	0.0977	0.9867	0.7548	0.7377	COVID-19 in healthcare workers in Italy: a prospective case-	
	0.0330	0.6270	0.0340	0.6180	-0.0001	1.0079	0.0328	0.6253	0.0342	0.6197	control study. Journal of Clinical Medicine, 9(9), 2870.	
					-0.0002		0.3661	0.7463	0.3675	0.7473		
					-0.0001		0.4666	0.7597	0.4379	0.7811		
					0.0001		0.4683		0.7360	0.8132		
					0.0000		0.3000		0.3000	0.7200	•	
					0.1687		_	0.9200	0.0108	0.0526		
					0.0001		_	0.9341	0.1645	0.7033		
					0.2175			0.9198	0.0178	0.0567		
					-0.0372		0.0216	0.4750	0.5375	0.6869		
					0.0000		0.4002		0.6668	0.7929		
					0.0000 0.0251		0.3668	0.9950	0.7862	0.7912 0.2322	1	
					0.0000		#DIV/0!			0.0000		
					-0.0001		0.5656		0.7401	0.8417		
		0.9730				1.0004	0.4025	0.9733	0.8557	0.8004		
12		0.7196				1.0000	0.9501	0.7200	0.3873	0.9872	Liu S, Yao N, Qiu Y & He C. (2020). Predictive performance of	Table 4
		0.8318				1.0000	0.8999	0.8316	0.5003	0.9780	SOFA and qSOFA for in-hospital mortality in severe novel	
	0.7000	0.8785	0.5185	0.9400	0.0000	1.0000	0.7000	0.8785	0.5185	0.9400	coronavirus disease. The American Journal of Emergency	
	0.5500	0.9439	0.6471	0.9182	0.0000	1.0000	0.5502	0.9439	0.6469	0.9181	Medicine, 38(10), 2074-2080.	
	0.2500	0.9813	0.7143	0.8750	0.0000	1.0000	0.2501	0.9813	0.7142	0.8749	· •	
	0.2000	0.9813	0.6667	0.8678	0.0000	1.0001	0.2001	0.9813	0.6665	0.8677		
	0.7000	0.8037	0.4000	0.9348	0.0000	1.0001	0.7001	0.8038	0.3999	0.9348		
	1.0000	0.4583	0.3953	1.0000	0.0000	1.0000	1.0000	#DIV/0!	#DIV/0!	1.0000		

#		Origina	l values			cking ues		Compute	ed values	1	Source	Data location
	Sens <sub>ij</sub>	Spec <sub>ij</sub>	PPV <sub>ij</sub>	NPV <sub>ij</sub>	DCD <sub>ij</sub>	DCR <sub>ij</sub>	Sens <sub>ij</sub>	Spec <sub>ij</sub>	PPV <sub>ij</sub>	NPV <sub>ij</sub>		in Source
	0.9412	0.7292	0.5517	0.9722	0.0000	1.0000	0.9411	0.7289	0.5521	0.9722		
		0.8542				1.0000	0.8823	0.8541	0.6819	0.9535		
					0.0000	1.0000	0.7059	0.8958	0.7059	0.8958		
					-0.0001		0.5871		0.8339	0.8684		
						0.9999	0.2350	0.9792	0.8003	0.7836		
						0.9998	0.1762	0.9792	0.7503	0.7708		
						1.0000	#DIV/0!	1.0000		#DIV/0!		
					0.0000	1.0000	0.7058	0.8542		0.8913		
13					-0.0002		0.6576		0.5347	0.9376	Covino M, Sandroni C, Santoro M, Sabia L, Simeoni B, Bocci MG	Table 2
					0.0001		0.7008	0.6488	0.2583	0.9248	& Franceschi, F. (2020). Predicting intensive care unit admission	
					-0.0001		0.6587	0.8483	0.4354	0.9344	and death for COVID-19 patients in the emergency department	
		0.7850				1.0005	0.7007	0.7856	0.3642	0.9368	using early warning scores. Resuscitation, 156, 84-91.	
		0.8270				1.0048	0.6278	0.8317	0.4698	0.8949		
		0.7960				1.0042	0.3410	0.7967	0.2263	0.8725		
					-0.0001		0.5995	0.7496	0.2975	0.9142		
					-0.0002		0.5868	0.8952	0.5342	0.9157		
		0.6480				1.0006 1.0008	0.6614 0.7153	0.6484	0.2737 0.3865	0.9049 0.9306		
					0.0001	1.0000	0.7153		0.3959	0.9300		
					0.0000	1.0050	0.6095	0.7910	0.5342	0.8711		
						0.9994	0.0095	0.7949	0.3342	0.8531		
		0.7550				1.0004	0.5893	0.7552		0.8331		
14	0.7895					0.9999	0.7894	0.7332	0.3751	0.9420	Hui TC, Khoo HW, Young BE, Mohideen SMH, Lee YS, Lim J &	Table 4
17					0.0000	0.9999	0.7034		0.2251	0.9710	Tan CH. (2020). Clinical utility of chest radiography for severe	Table +
		0.9111				1.0000	0.6315		0.6001	0.9213	COVID-19. Quantitative imaging in medicine and surgery, 10(7),	
		0.8776				1.0000	0.7272		0.4001	0.9663	1540.	
		0.6364				1.0000	0.9444	0.6365	0.5151	0.9655		Table 5
		0.9545				1.0000	0.8889		0.8889	0.9545		1 4510 0
					0.0000	1.0000	1.0000		#DIV/0!	_		
15	0.9390					0.9999	0.9378		0.9491	0.9530	Gezer NS, Ergan B, Barış MM, Appak Ö, Sayıner AA, Balcı P &	
						1.0000	0.9900		0.8586	0.9910	Kılınç, O. (2020). COVID-19 S: A new proposal for diagnosis and	
											structured reporting of COVID-19 on computed tomography	
											imaging. Diagnostic and Interventional Radiology, 26(4), 315.	
16	0.7200	0.9400	0.9200	0.7600	-0.0012	0.9973	0.6992	0.9340	0.9271	0.7779	Bai, H.X, Hsieh B, Xiong Z, Halsey K, Choi JW, Tran TML &	Table 3
	0.7200	0.8800	0.8700	0.7400	0.0002	1.0006	0.7220	0.8811	0.8689		Jiang, X. L. (2020). Performance of radiologists in differentiating	

#		Origina	l values			cking ues		Compute	ed values	5	Source	Data location
	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	$NPV_{ij}$	$DCD_{ij}$	DCR <sub>ij</sub>	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	$NPV_{ij}$		in Source
					0.0002 0.0000	1.0020 1.0000	0.9404 #DIV/0!	0.2414 1.0000	0.5681 1.0000	0.7887 #DIV/0!	COVID-19 from viral pneumonia on chest CT. Radiology, 200823.	Table 4
					-0.0006		0.6618	0.9276	0.9130	0.7274		
					-0.0006		0.9682	0.0661	0.5452	0.6834		
		0.9300				1.0000 1.0011	#DIV/0! 0.8400	1.0000 0.9346	1.0000 0.9251	#DIV/0! 0.8300		
		0.9300				1.0011	0.7327	0.9346	0.9251	0.6300		
					0.0002	1.0004	#DIV/0!	1.0000	1.0000	#DIV/0!		
17	0.8300					1.3929	0.9574	0.7383	0.2700	0.6373	Tan C, Huang Y, Shi F, Tan K, Ma Q, Chen Y, Jiang X, Li X.	Table 2
	0.8300	0.4300	0.2900	0.9000	0.0000	0.9993	0.8297	0.4295	0.2904	0.9002	C-reactive protein correlates with computed tomographic findings	
					0.0000			0.6700	0.2200	0.7800	and predicts severe COVID-19 early. Journal of Medical Virology.	
						0.9993	0.8297	0.4295	0.2904	0.9002	2020;92(7):856-62.	
					-0.0006		0.8214	0.9050	0.7221	0.9527		
					-0.0007		0.8239	0.8033	0.5706	0.9424		
18					-0.0023 0.0000	1.0000	0.4768 0.6000	0.9021 0.5400	0.6221	0.8708 0.5400	Caruso D, Zerunian M, Polici M, Pucciarelli F, Polidori T, Rucci C	Table 2
10		0.3400				1.0000	0.9098		0.4811		& Laghi, A. (2020). Chest CT features of COVID-19 in Rome,	I able 2
	0.0000	0.1000	0.0001	0.0700	0.0000	1.0010	0.0000	0.0110	0.4011	0.0011	Italy. Radiology, 201237.	
19	0.8630	0.5930	0.3390	0.9474	0.0001	1.0008	0.8638	0.5946	0.3376	0.9471	Bi X, SU Z, Yan H, Du J, Wang J, Chen L & Li J. (2020).	Table
	0.8570	0.4290	0.9000	0.3330	-0.0001	0.9994	0.8567	0.4285	0.9002	0.3335	Prediction of severe illness due to COVID-19 based on an	3.3
											analysis of initial Fibrinogen to Albumin Ratio and Platelet count.	
	0.0070	0.0000	0.0440	0.0000	0.0000	1.0000	0.0000	0.0000	0.0440	0.0000	Platelets, 1-6.	
20	0.8670				0.0000	1.0000	0.8669 0.9559		0.9110 0.9153	0.9030 0.9651	Dangis A, Gieraerts C, Bruecker YD, Janssen L, Valgaeren H,	Table 2
	0.9560	0.9320	0.9150	0.9650	0.0000	1.0000	0.9559	0.9318	0.9153	0.9651	Obbels D & Symons R. (2020). Accuracy and reproducibility of low-dose submillisievert chest CT for the diagnosis of COVID-19.	
											Radiology: Cardiothoracic Imaging, 2(2), e200196.	
21	0.5200	0.8900	0.2600	0.9600	-0.0005	0.9949	0.5103	0.8862	0.2675	0.9615	Laguna-Goya R, Utrero-Rico A, Talayero P, Lasa-Lazaro M,	
	0.9700	0.5300	0.1400	0.9900	-0.0025	0.9644	0.9346	0.3326	0.2692	0.9956	Ramirez-Fernandez A, Naranjo L & Fernández-Ruiz M. (2020). IL-	
					-0.0049		0.6513	0.6994		0.9801	6-based mortality risk model for hospitalized patients with	
					0.0002		0.8021		0.1582		COVID-19. Journal of Allergy and Clinical Immunology, 146(4),	
					-0.0003	_	0.7155	0.7055	0.1630	0.9706	799-807.	
					-0.0034 -0.0026		0.6370	0.6929	0.1780	0.9773		
					-0.0026		0.6002	_	0.2630	0.9663 0.9556		
					-0.0020	= " " " " " " " " " " " " " " " " " " "	0.5937		0.1229	0.9330		
	0.1100	0.0700	0.1100	0.000	0.0104	0.1112	0.0001	0.7000	0.2201	0.3021		

#		Origina	l values			cking lues		Compute	ed values	3	Source	Data location
	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	NPV <sub>ij</sub>	DCD <sub>ij</sub>	DCR <sub>ij</sub>	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	$NPV_{ij}$		in Source
					-0.0012		0.8413	0.5847	0.1698	0.9827		
					-0.0011		0.5999			0.9823		
					-0.0022		0.8253	0.6203	_	0.9846		
					0.0004		0.8849		0.3099	0.9791		
					0.0004		0.7587	0.8849	0.3099	0.9791		
					0.0012		0.8823	0.8563		0.9878		
					-0.0018		0.7749	0.8555		0.9838		
					-0.0018		0.7700	0.9003	0.5001			
					0.0001	1.0004	0.8823		0.3747			
22	0.7690					0.9999	0.7689	0.7259	0.4472		Jehi L, Ji X, Milinovich A, Erzurum S, Merlino A, Gordon S &	Table 2
						1.0209	0.5841	0.9358		0.8688	Kattan, MW. (2020). Development and validation of a model for	
					-0.0001		0.3864	0.9628	0.7502		individualized prediction of hospitalization risk in 4,536 patients	
					-0.0002		0.2486	0.9785	0.7761	0.8234	with COVID-19. PloS one, 15(8), e0237419.	
23	0.1170				-0.0002	1.0015	0.1118 0.5015	0.9916 0.8288	0.8081	0.8042 0.9135	Crustay IA Dyambinatiy A Kannaria M.A. LaDrun DC	Toble 2
23		0.8280				1.0015	0.5015	0.8288	0.5000	0.9135	Gruskay JA, Dvorzhinskiy A, Konnaris M.A, LeBrun DG,	Table 2
					0.0000	1.0006	0.3000		0.5000	0.9220	Ghahramani GC, Premkumar A & Ricci WM. (2020). Universal testing for COVID-19 in essential orthopaedic surgery reveals a	
					0.0002			0.7671		0.9111	high percentage of asymptomatic infections. JBJS, 102(16), 1379-	
											1388.	
24					-0.0005		0.8876	0.6545	0.6357	0.9022	Gidari A, De Socio G V, Sabbatini S & Francisci D. (2020).	Table 2
	0.6300	0.9800	0.9400	0.8000	-0.0018	0.9960	0.5612	0.9735	0.9543	0.8419	Predictive value of National Early Warning Score 2 (NEWS2) for	
											intensive care unit admission in patients with SARS-CoV-2	
											infection. Infectious Diseases, 52(10), 698-704.	
25					0.0013	1.0031	0.6883	0.9353	0.7862	0.8709	Himoto Y, Sakata A, Kirita M, Hiroi T, Kobayashi KI, Kubo K, Kim	Table 3
		0.6000				1.0009	0.8308	0.6013	0.4486	0.8995	H, Nishimoto A, Maeda C, Kawamura A, Komiya N. Diagnostic	
					0.0002	1.0009	0.8308		0.4486	0.8995	performance of chest CT to differentiate COVID-19 pneumonia in	
		0.6700				1.0022	0.8328	0.6744	0.4950	0.9084	non-high-epidemic area in Japan. Japanese Journal of Radiology.	
		0.6000				1.0009	0.8308	0.6013	0.4486	0.8995	2020;38(5):400-406.	
-00		0.8000				1.0001	0.8304	0.8004	0.6294	0.9198	Luc V Maa L Vusa V Vus V Lis O Tara C 9 Cus 7 (0000)	Table 0
26		0.9215				0.9999	0.4117		0.2801	0.9548	Luo Y, Mao L, Yuan X, Xue Y, Lin Q, Tang G & Sun Z. (2020).	Table 2
		0.9041 0.9026				0.9999	0.6860 0.5491	0.9040 0.9026	0.3468 0.2946	0.9749 0.9643	Prediction model based on the combination of cytokines and lymphocyte subsets for prognosis of SARS-CoV-2 infection.	
		0.9020				0.9999	0.3332	0.9026		0.9643	Journal of Clinical Immunology, 40(7), 960-969.	
		0.9070				1.0000	0.3332	0.9072		0.9527	oduriai di Ciinicai inimundidgy, 40(1), 300-303.	
					0.0000	1.0000	0.3922		0.2301	0.9327		
	0.7043	0.9041	0.5774	0.9020	0.0000	1.0000	0.7041	0.9040	0.5111	0.3020		

#		Origina	l values			cking ues		Compute	ed values	•	Source	Data location
	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	$NPV_{ij}$	DCD <sub>ij</sub>	DCR <sub>ij</sub>	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	$NPV_{ij}$		in Source
	0.7255	0.9099	0.3737	0.9781	0.0000	0.9999	0.7252	0.9098	0.3741	0.9781		
	0.6667	0.9055	0.3434	0.9734	0.0000	0.9999	0.6664	0.9054	0.3437	0.9734		
	0.9020	0.9026	0.4071	0.9920	0.0000	1.0000	0.9018	0.9024	0.4075	0.9920		
27	0.8934	0.9016	0.9008	0.8943	0.0000	1.0000	0.8934	0.9016	0.9008	0.8943	Ardakani AA, Acharya UR, Habibollahi & Mohammadi, A. (2021).	Table 2
	0.7903	0.8710	0.8596	0.8060	0.0000	1.0000	0.7902	0.8710	0.8596	0.8061	COVIDiag: A clinical CAD system to diagnose COVID-19	
	0.9098	0.8975	0.8987	0.9087	0.0000	1.0000	0.9098	0.8975	0.8987	0.9087	pneumonia based on CT findings. European radiology, 31(1),	
	0.8871	0.8064	0.8209	0.8772	0.0000	1.0000	0.8871	0.8065	0.8208	0.8772	121-130.	
	0.8852	0.9344	0.9310	0.8906	0.0000				0.9310	0.8906		
	0.8548	0.8871	0.8833	0.8594	0.0000	1.0000	0.8548	0.8871	0.8833	0.8594		
				0.9336		1.0000	0.9345		0.9230	0.9335		
				0.8852		1.0000		0.8709	0.8731	0.8853		
				0.9458		1.0000	0.9468	0.9305	0.9314	0.9457		
				0.9333		1.0000	0.9355	0.9034	0.9062	0.9332		
					-0.0003		0.8678		0.8853	0.8761		
28					-0.0005		0.8632	0.7904	0.3533	0.9811	Larner AJ. (2021). Cognitive testing in the COVID-19 era: can	Table 3
					-0.0017		0.7729			0.9744	existing screeners be adapted for telephone use?.	
					0.0012		0.8449		0.4808	0.8555	Neurodegenerative Disease Management, 11(1), 77-82.	
					-0.0010				0.6662	0.7749		
29					-0.0012		0.9342		0.8248	0.9086	Rueckel J, Fink N, Kaestle S, Stüber T, Schwarze V, Gresser E,	Table 2
					0.0004		0.7300		0.9780	0.7100	Hoppe BF, Rudolph J, Kunz WG, Ricke J, Sabel BO. COVID-19	
					0.0014				0.7833	0.8905	Pandemic and Upcoming Influenza Season—Does an Expert's	
	0.7400	0.9800	0.9800	0.7500	0.0002	1.0004	0.7500	0.9810	0.9789	0.7400	Computed Tomography Assessment Differentially Identify COVID-	
											19, Influenza and Pneumonias of Other Origin?. Journal of	
	0.0000	0.0050	0.0440	0.5500	0.0000	0.0004	0.0074	0.0040	0.0440	0.5547	Clinical Medicine. 2021;10(84):1-13.	T-61-
30							0.6274		0.9446	0.5547	Liu J, Lian R, Zhang G, Hou B, Wang C, Dong J & Ye, T. (2021).	Table
					-0.0001		0.7760	0.9494	0.9693	0.6805	Changes in serum virus-specific IgM/IgG antibody in	S1
					-0.0001		0.8514		0.9453	0.7510	asymptomatic and discharged patients with reoccurring positive	
	0.9510	0.7500	0.8850	0.8850	0.0002	1.0003	0.9518	0.7532	0.8833	0.8833	COVID-19 nucleic acid test (RPNAT). Annals of medicine, 53(1),	
-04	0.7500	0.0040	0.0040	0.0050	0.0004	0.0000	0.7504	0.0000	0.0005	0.0050	34-42.	Table 0
31					-0.0001		0.7564	0.8623	0.2335	0.9852	Spangler D, Blomberg H Smekal D. (2021). Prehospital	Table 3
				0.9880		1.0007	0.8236		0.1675	0.9879	identification of Covid-19: an observational study. Scandinavian	
					0.0000		0.7110		0.2470	0.9530	Journal of Trauma, Resuscitation and Emergency Medicine,	
					-0.0002		0.7890		0.1788	0.9545	29(1), 1-10.	
32				1.0000		1.0000	1.0000		#DIV/0!		Akhavan AR, Habboushe JP, Gulati R, Iheagwara O, Watterson J,	Table 3
	0.9800	0.1500	0.4100	0.9400	0.0013	1.0269	0.9840	0.1818	0.3556	0.9256	Thomas S, Swartz JL, Koziatek CA, Lee DC. Risk Stratification of	

#		Origina	l values			king ues		Compute	ed values	3	Source	Data location
	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	$NPV_{ij}$	DCD <sub>ij</sub>	DCR <sub>ij</sub>	Sens <sub>ij</sub>	Spec <sub>ij</sub>	PPV <sub>ij</sub>	$NPV_{ij}$		in Source
	0.9700	0.1800	0.4200	0.9000	-0.0008	0.9854	0.9674	0.1678	0.4409	0.9074	COVID-19 Patients Using Ambulatory Oxygen Saturation in the	
	0.9200	0.2900	0.4400	0.8600	0.0006	1.0072	0.9220	0.2956	0.4333	0.8567	Emergency Department. Western Journal of Emergency Medicine. 2020;21(6):5-14.	
	0.8900	0.3600	0.4500	0.8400	-0.0016	0.9832	0.8842	0.3468	0.4643	0.8476	Medicine. 2020,21(0):3-14.	
	0.8100	0.4300	0.4600	0.7900	-0.0001	0.9984	0.8094	0.4291	0.4609	0.7906		
	0.7300	0.5000	0.4700	0.7500	-0.0008	0.9904	0.7268	0.4960	0.4740	0.7530		
	0.6000	0.5900	0.4700	0.7100	0.0003	1.0050	0.6014	0.5914	0.4686	0.7088		
	0.3500	0.8000	0.5100	0.6700	-0.0009	0.9833	0.3457	0.7969	0.5148	0.6742		
	0.3400	0.8500	0.5800	0.6800	0.0002	1.0027	0.3412	0.8507	0.5787	0.6789		
	0.9800	0.0500	0.1000	0.9600	0.0001	1.0208	0.9806	0.0516	0.0970	0.9587		
	0.9800	0.1100	0.1000	0.9800	-0.0002	0.9778	0.9778	0.1000	0.1100	0.9820		
	0.9000	0.2300	0.1100	0.9500	-0.0012	0.9142	0.8872	0.2069	0.1240	0.9560		
	0.8200	0.4200	0.1300	0.9600	0.0010	1.0349	0.8320	0.4405	0.1208	0.9567		
	0.7400	0.5800	0.1600	0.9500	-0.0014	0.9706	0.7238	0.5598	0.1714	0.9538		
	0.5400	0.7300	0.1800	0.9400	0.0016	1.0355	0.5599	0.7455	0.1685	0.9353		
	0.2800	0.8300	0.1500	0.9200	0.0011	1.0717	0.2936	0.8392	0.1417	0.9150		
	0.2000	0.9000	0.1700	0.9100	-0.0011	0.9308	0.1871	0.8923	0.1820	0.9166		
	0.0800				-0.0004		0.0743	0.9568	0.1711	0.9164		
	0.0200				0.0000			0.9800	0.1000	0.9000		
33					-0.0002		0.9885	0.1996	0.4944	0.9617	Sung J, Choudry N, Bachour R. Development and validation of a	Table 5
					0.0000		0.9771	0.4747	0.5813	0.9649	simple risk score for diagnosing COVID-19 in the emergency	
					-0.0002		0.9795	0.1593	0.4066	0.9347	room. Epidemiology & Infection. 2020;148(e273):1-7.	
	0.2760				-0.0001		0.2727	0.9827	0.9242	0.6478		
	0.2860			0.7060	-0.0001		0.2835	0.9767	0.8763	0.7085		
34	0.9630			0.9970		1.0005	0.9638	0.6581	0.1284	0.9969	Dagan N, Barda N, Riesel D, Grotto I, Sadetzki S & Balicer R.	
					-0.0003		0.9124		0.2099		(2020). A score-based risk model for predicting severe COVID-19	
						1.0007	0.4326	0.9525	0.3237	0.9687	infection as a key component of lockdown exit strategy. medRxiv.	Table 2
					0.0000 0.0322		0.4886 0.5038	0.8428 0.7482	0.1452	0.9680 0.2626	Available at: https://www.medrxiv.org/content/medrxiv/early/2020/05/23/2020.0	
					0.0002		0.9666	0.6009	0.1239	0.9955	5.20.20108571.full.pdf.	
					-0.0002		0.9148		0.2049	0.9924	· r·	

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#	Original values					cking ues	Computed values				Source	Data location	
	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	NPV <sub>ij</sub>	DCD <sub>ij</sub>	DCR <sub>ij</sub>	Sens <sub>ij</sub>	Spec <sub>ij</sub>	$PPV_{ij}$	$NPV_{ij}$		in Source	
	0.4300	0.9380	0.3260	0.9590	-0.0001	0.9992	0.4278	0.9375	0.3279	0.9593			
	0.4890	0.8000	0.1450	0.9570	-0.0002	0.9950	0.4855	0.7977	0.1468	0.9576			
	0.0810	0.2620	0.0800	0.8040	0.0398	0.0583	0.5012	0.8019	0.0076	0.2646			
	0.5430	0.9450	0.2130	0.9870	0.0000	1.0003	0.5446	0.9453	0.2119	0.9869			
	0.4570	0.8340	0.0700	0.9820	-0.0002	0.9908	0.4497	0.8299	0.0719	0.9825			
	0.5430	0.9300	0.2130	0.9830	-0.0001	0.9994	0.5409	0.9294	0.2145	0.9831			
	0.4570	0.7890	0.0700	0.9770	0.0001	1.0073	0.4609	0.7916	0.0690	0.9766			
35	0.7780	0.8310	0.3220	0.9730	-0.0001	0.9996	0.7768	0.8300	0.3235	0.9732	Ma J Shi X, Xu W, Lv F, Wu J, Pan Q, Yang J, Yu J, Cao H, Li L.	Table 4	
	0.7720	0.8180	0.2590	0.9780	0.0002	1.0014	0.7756	0.8211	0.2550	0.9775	Development and validation of a risk stratification model for screening suspected cases of COVID-19 in China. Aging (Albany NY). 2020;12(14):13882-13894.		

	RED	ORANGE		YELLOW		WHITE		YELLOW		ORANGE		RED
DCR=1	< 0.9400	0.9400	0.9599	0.9600	0.9799	0.9800	1.0199	1.0200	1.0399	1.0400	1.0599	>= 1.06
DCD=0	< -0.0600	-0.0600	-0.0410	-0.0400	-0.021	-0.0200	0.0199	0.0200	0.0399	0.0400	0.0599	>= 0.06
Others ,%	<= -6%	- 5.999% to - 4%		- 3.999% to - 2.0%		- 2.0% to 2.0%		2% to 3.999%		4% to 5.999%		>= 6%

### 5. CONCLUSIONS

We provide a novel method for validating any purported set of the four most prominent indicators of diagnostic testing (Sensitivity, Specificity, Positive Predictive Value, and Negative Predictive Value), by observing that these indicators constitute three rather than four independent quantities. This observation has virtually been unheard of in the open medical literature. Up to date, researchers who were unaware of this observation could possibly be somewhat excused. but, from now on, this observation should be popularized and respected. We defined two functions that check consistency for any set of four numerical values claimed to be the four basic diagnostic indicators, and naturally belonging to the interval [0.0, 1.0]. Most of data we came across herein and earlier in [17,45] met our criterion for consistency, but in a few cases, there were obvious blunders. The percentage of inconsistent cases and the severity of deviations therein seem definitely higher for the present COVID-19 data than they were in [17,45]. Our preliminary results seem to strengthen our suspicions of compromising the quality of publications in the COVID-19 era.

Our study has a few limitations that allow for possible improvements in future work. Our validating formulas express any one of the four diagnostic indicators in terms of the other three, only under the assumptions that each of the four indicators exists, and that no division by zero is encountered. Though our validating formulas are correct for any combination of input values, they occasionally produce the indefinite value (0/0), which need to be defined through a limit operation [4]. The threshold for accepting or rejecting purported sets is selected in an arbitrary (albeit plausible) way. Such a threshold might be chosen in other (statistically-rigorous) ways. Another major limitation of our work is that our coverage of publications in the COVID-19 era is far from being exhaustive. We did not set up an organized plan to cover all publications dealing with COVID-19, or to select a representative sample of these publications. We simply presented an adequate number of publications that we came across, and that we could access, in the Google Scholar database, in which we could locate a complete set of the four diagnostic indicators reported on COVID-19. We honestly report the results of all these publications that we have seen, with no exclusion of any of them. In summary, we suggest that some further literature-encompassing statisticallyrigorous investigation of the pandemic-induced quality deterioration should be pursued.

### **COMPETING INTERESTS**

The authors have declared that no competing interests exist.

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   Available:https://doi.org/10.3390/ani71200 90

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