Random Fourier Signature Features

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Abstract

Tensor algebras give rise to one of the most powerful measures of similarity for sequences of arbitrary length called the signature kernel accompanied with attractive theoretical guarantees from stochastic analysis. Previous algorithms to compute the signature kernel scale quadratically in terms of the length and the number of the sequences. To mitigate this severe computational bottleneck, we develop a random Fourier feature-based acceleration of the signature kernel acting on the inherently non-Euclidean domain of sequences. We show uniform approximation guarantees for the proposed unbiased estimator of the signature kernel, while keeping its computation linear in the sequence length and number. In addition, combined with recent advances on tensor projections, we derive two even more scalable time series features with favourable concentration properties and computational complexity both in time and memory. Our empirical results show that the reduction in computational cost comes at a negligible price in terms of accuracy on moderate-sized datasets, and it enables one to scale to large datasets up to a million time series.

1 Introduction

Machine learning has successfully been applied to tasks that require learning from complex and structured data types on non-Euclidean domains. Feature engineering on such domains is often tackled by exploiting the geometric structure and symmetries existing within the data [5]. Learning from sequential data (such as video, text, audio, time series, health data, etc.) is a classic, but an ongoing challenge due to the following properties:

- *Non-Euclidean data*. The data domain is nonlinear since there is no obvious and natural way of adding sequences of different length.
- *Time-space patterns*. Statistically significant patterns can be distributed over time and space, that is, capturing the order structure in which "events" arise is crucial.
- *Time-warping invariance.* The meaning of many sequences is often invariant to reparametrization also frequently called time-warping, at least to an extent; e.g. a sentence spoken quicker or slower contains (essentially) the same information.

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- *Discretization and irregular sampling.* Sequences often arise by sampling along an irregularly spaced grid of an underlying continuous time process. A general methodology should be robust as the sampling gets finer, sequences approximate paths (continuous-time limit), or as the discretization grid varies between sequences.
- *Scalability.* Sequence datasets can quickly become massive, so the computational complexity should grow subquadratically, in terms of all of the state-space dimension, and the length and number of sequences.

The signature kernel k_{Sig} is the state-of-the-art kernel for sequential data [84, 68, 47] that addresses the first 4 of the above questions and can rely on the modular and powerful framework of kernel learning [71]. Its construction is motivated by classic ideas from stochastic analysis that give a structured description of a sequence by developing it into a series of tensors. We refer to [46] for a recent overview of its various constructions and applications.

Features vs Kernel/Primal vs Dual. Kernel learning circumvents the costly evaluation of a high- or infinite-dimensional feature map by replacing it with the computation of a Gram matrix which contains as entries the inner products of features between all pairs of data points. This can be very powerful since the inner product evaluation can often be done cheaply by the celebrated "kernel trick", even for infinite-dimensional feature spaces, but the price is that now the computational cost is quadratic in the number of samples, and downstream algorithms further often incur a cubic cost usually in the form of a matrix inversion. On the other hand, when finite-dimensional features can used for learning, the primal formulation of a learning algorithm can perform training and inference in a cost that is linear with respect to the sample size assuming that the feature dimension is fixed. This motivates the investigation of finite-dimensional approximations to kernels that mimic their expressiveness at a lower computational cost. It is an interesting question how the feature dimension should scale with the dataset size to maintain a given (optimal) learning performance in downstream tasks, which is investigated for instance by [65, 7, 77, 48, 75, 43].

Computational Cost of the Signature kernel. In the context of the signature kernel, one data point is itself a whole sequence. Hence, given a data set **X** consisting of $N \in \mathbb{Z}_+$ sequences where each sequence $\mathbf{x} = (\mathbf{x}_1, \dots, \mathbf{x}_\ell)$ is of maximal length $\ell \leq L \in \mathbb{Z}_+$ and has sequence entries \mathbf{x}_i in a state-space of dimension d, then the existing algorithms to evaluate the Gram matrix of the k_{Sig} scale quadratically, i.e. as $O(N^2L^2d)$, both in the sequence length L and the number of sequences N. So far this has only been addressed by subsampling (either directly the sequence elements to reduce the length or by column subsampling via the Nyström approach [91]), which can lead to crude approximations and performance degradation on large-scale datasets.

Contribution. Random Fourier Features (RFF) [58] is a classic technique to enjoy both the benefits of the primal and dual approach. Here, a low-dimensional and random feature map is constructed, which although does not approximate the feature map of a translation-invariant kernel, its inner product is with high probability close to the kernel itself. The main contribution of this article is to carry out such a construction for the signature kernel. Concretely, we construct a random feature map on the domain of sequences called Random Fourier Signature Features (RFSF), such that its inner product is a random kernel \tilde{k}_{Sig} for sequences that is both (i) an unbiased estimator for k_{Sig} , and (ii) has analogous probabilistic approximation guarantees to the classic RFF kernel. The challenge is that a direct

application of the classic RFF technique is not feasible since this relies on Bochner's theorem which does not apply since the sequence domain is not even a linear space and the feature domain is non-Abelian, which makes the use of (generalizations of [28]) Bochner's theorem difficult due to the lack of sufficiently explicit representations. We tackle this challenge by combining the algebraic structure of signatures with probabilistic concentration arguments; a careful analysis of the error propagation yields uniform concentration guarantees similar to the RFF on \mathbb{R}^d . Then, we introduce dimensionality reduction techniques for random tensors further approximating \tilde{k}_{Sig} to define the extremely scalable variants $\tilde{k}_{\text{Sig}}^{\text{DP}}$ and $\tilde{k}_{\text{Sig}}^{\text{TRP}}$ called RFSF-DP and RFSF-TRP saving considerable amounts of computation time and memory by low-dimensional projection of the feature set of the RFSF.

Hence, analogously to the classic RFF construction, the random kernels \tilde{k}_{Sig} , \tilde{k}_{Sig}^{DP} , \tilde{k}_{Sig}^{TRP} simultaneously enjoy the expressivity of an infinite-dimensional feature space as well as linear complexity in sequence length. This overcomes the arguably biggest drawback of the signature kernel, which is the quadratic complexity in sample size and sequence length; the price for reducing the complexities by an order is that this approximation only holds with high probability. As in the case of the classic RFF, our experiments show that this is in general a very attractive tradeoff. Concretely, we demonstrate in the experiments that the proposed random features (1) provide comparable performance on moderate sized datasets to full-rank (quadratic time) signature kernels, (2) outperform other random feature approaches for time series on both small- and large-scale datasets, (3) allow scaling to datasets consisting of a million time series.

Related Work. The signature kernel has found many applications; for example, it is used in ABC-Bayes [22], economic scenario validation [1], amortised likelihood estimation [23], the analysis of RNNs [25], analysis of trajectories in Lie groups [45], metrics for generative modelling [6, 37], or dynamic analysis of topological structures [29]. For a general overview see [46]. All of these applications can benefit from a faster computation of the signature kernel with theoretical guarantees. Previous approaches address the quadratic complexity of the signature kernel only by subsampling in one form or another: [39] combine a structured Nyström type-low rank approximation to reduce complexity in dimension of samples and sequence length, [84] combine this with inducing point and variational methods, [68] uses sequence-subsampling, [47] use diagonal approximations to Gram matrices in a variational setting. Related to this work is also the random nonlinear projections in [52]; further, [54] combine linear dimension projection in a general pipeline and [15] use signatures in reservoir computing. Directly relevant for this work is recent progress on tensorized random projections [78, 61]. Random Fourier Features [58, 60] are well-understood theoretically [79, 74, 75, 49, 2, 80, 9, 85, 9]. In particular, its generalization properties are studied in e.g. [3, 48, 77, 43], where it is shown that the feature dimension need only scale sublinearly in the dataset size for supervised learning, and a similar result also holds for kernel principal component analysis [75]. Several variations have been proposed over the years [44, 24, 13, 94, 12, 11, 14], even finding applications in deep learning [81]. Alternative random feature approaches for polynomial and Gaussian kernels based on tensor sketching have been proposed in e.g. [90, 89, 88]. For a survey, the reader is referred to [9, 50].

Outline. Section 2 provides background on the prerequisites of our work: Random Fourier Features, and Signature Features/Kernels. Section 3 contains our proposed methods with theoretical results; it introduces Random Fourier Signature Features (RFSF) $\tilde{\varphi}_{\text{Sig}_{\leq M}}$, RFSF kernels $\tilde{k}_{\text{Sig}_{\leq M}}$ (where $M \in \mathbb{Z}_+$ is the truncation level introduced later), and most impor-

tantly their theoretical guarantees. Theorem 3.2 quantifies the approximation $k_{Sig_{\leq M}}(\mathbf{x}, \mathbf{y}) \approx \tilde{k}_{Sig_{\leq M}}(\mathbf{x}, \mathbf{y})$ uniformly. Then, we discuss additional variants: the RFSF-DP kernel $\tilde{k}_{Sig_{\leq M}}^{DP}$ and the RFSF-TRP kernel $\tilde{k}_{Sig_{\leq M}}^{TRP}$, which build on the previous construction using dimensionality reduction with corresponding concentration results in Theorems 3.5 and 3.8. Section 4 compares the performance of the proposed scalable signature kernels against popular approaches on SVM multivariate time series classification, which demonstrates that the proposed kernel not only significantly improves the computational complexity of the signature kernel, it also provides comparable performance, and in some cases even improvements in accuracy as well. Hence, we take the best of both worlds: linear batch, sequence, and state-space dimension complexities, while approximately enjoying the expressivity of an infinite-dimensional RKHS with high probability.

2 **Prerequisites**

Notation. We denote the real numbers by \mathbb{R} , natural numbers by $\mathbb{N} \coloneqq \{0, 1, 2, ...\}$, positive integers by $\mathbb{Z}_+ \coloneqq \{1, 2, 3, ...\}$, the range of positive integers from 1 to $n \in \mathbb{Z}_+$ by $[n] \coloneqq \{1, 2, ..., n\}$. Given $a, b \in \mathbb{R}$, we denote their maximum by $a \lor b \coloneqq \max(a, b)$ and their minimum by $a \land b \coloneqq \min(a, b)$. We define the collection of all ordered *m*-tuples with non-repeating entries starting from 1 up to *n* including the endpoints by

$$\Delta_m(n) \coloneqq \{ 1 \le i_1 < i_2 < \dots < i_m \le n : i_1, i_2, \dots, i_n \in [n] \}.$$
⁽¹⁾

In general, \mathcal{X} refers to a subset of the input domain, where the various objects are defined, generally taken to be a subset \mathbb{R}^d unless otherwise stated. For a vector $\mathbf{x} \in \mathbb{R}^d$, we denote its ℓ_p norm by $\|\mathbf{x}\|_p \coloneqq \left(\sum_{i=1}^p |x_i|^p\right)^{1/p}$. For a matrix $\mathbf{A} \in \mathbb{R}^{d \times e}$, we denote the spectral and the Frobenius norm by $\|\mathbf{A}\|_2 \coloneqq \sup_{\|\mathbf{x}\|_2=1} \|\mathbf{A}\mathbf{x}\|_2$ and $\|\mathbf{A}\|_F \coloneqq \left(\sum_{i=1}^e \|\mathbf{A}\mathbf{e}_i\|_2^2\right)^{1/2}$, where $\{\mathbf{e}_1, \dots, \mathbf{e}_d\}$ is the canonical basis of \mathbb{R}^d . The transpose of a matrix \mathbf{A} is denoted by \mathbf{A}^{\top} . For a differentiable $f : \mathbb{R}^d \to \mathbb{R}$, we denote its gradient at $\mathbf{x} \in \mathbb{R}^d$ by $\nabla f(\mathbf{x}) \coloneqq \left(\frac{\partial f(\mathbf{x})}{\partial x_i}\right)_{i=1}^d$, and its collection of partial derivatives with respect to $\mathbf{s} \coloneqq (x_{i_1}, \dots, x_{i_k})$ by $\partial_{\mathbf{s}} f(\mathbf{x}) \coloneqq \left(\frac{\partial f(\mathbf{x})}{\partial x_{i_j}}\right)_{j=1}^k$. \mathcal{X}_{seq} refers to sequences of finite, but unbounded length with values in the set \mathcal{X} :

$$\mathcal{X}_{seq} \coloneqq \{\mathbf{x} = (\mathbf{x}_1, \dots, \mathbf{x}_L) : \mathbf{x}_i \in \mathcal{X}, L \in \mathbb{Z}_+\}.$$

We denote the length of a sequence $\mathbf{x} = (\mathbf{x}_1, \dots, \mathbf{x}_L) \in \mathcal{X}_{seq}$ by $\ell_{\mathbf{x}} \coloneqq L$, and define the 1st-order forward differencing operator as $\delta \mathbf{x}_i \coloneqq \mathbf{x}_{i+1} - \mathbf{x}_i$. We define the 1-variation functional of a sequence $\mathbf{x} \in \mathcal{X}_{seq}$ as $\|\mathbf{x}\|_{1-var} \coloneqq \sum_{i=1}^{\ell_{\mathbf{x}}-1} \|\delta \mathbf{x}_i\|_2$ as a measure of sequence complexity.

Random Fourier Features. Kernel methods allow to implicitly use an infinite-dimensional feature map $\varphi : \mathfrak{X} \to \mathcal{H}$ by evaluation of the inner product $k(\mathbf{x}, \mathbf{y}) = \langle \varphi(\mathbf{x}), \varphi(\mathbf{y}) \rangle_{\mathcal{H}}$, when \mathcal{H} is a Hilbert space. This inner product can often be evaluated without direct computation of $\varphi(\mathbf{x})$ and $\varphi(\mathbf{y})$ via the kernel trick. Although this makes them a powerful tool due to the resulting flexibility, the price of this flexibility is a trade-off in complexity with respect to the number of samples $N \in \mathbb{Z}_+$. Disregarding the price of evaluating the kernel $k(\mathbf{x}, \mathbf{y})$ momentarily, kernel methods require the computation of a Gram matrix with $O(N^2)$ entries, that further incurs an $O(N^3)$ computational cost by most downstream algorithms, such as KRR [72], GP [63], and SVM [71]. Several techniques reduce this complexity, and the focal point

of this article is the Random Fourier Feature (RFF) technique of [58, 59, 60], which can be applied to any continuous, bounded, translation-invariant kernel on \mathbb{R}^{d} .¹ Throughout, we write with some abuse of notation $k(\mathbf{x}-\mathbf{y}) \equiv k(\mathbf{x}, \mathbf{y})$. Next, we outline the RFF construction.

A corollary of Bochner's theorem [67] is that any continuous, bounded, and translationinvariant kernel $k : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ can be represented as the Fourier transform of a nonnegative finite measure Λ called the spectral measure associated to k, i.e. for $\mathbf{x}, \mathbf{y} \in \mathcal{X}$

$$\mathbf{k}(\mathbf{x}-\mathbf{y}) = \int_{\mathbb{R}^d} \exp(i\mathbf{w}^{\mathsf{T}}(\mathbf{x}-\mathbf{y})) \mathrm{d}\Lambda(\mathbf{w}).$$

We may, without loss of generality, assume that Λ is a probability measure such that $\Lambda(\mathbb{R}^d) = 0$, which amounts to working with the kernel $k(\mathbf{x}-\mathbf{y})/k(\mathbf{0})$. [58] proposed to draw $\tilde{d} \in \mathbb{Z}_+$ i.i.d. samples from Λ , $\mathbf{w}_1, \ldots, \mathbf{w}_{\tilde{d}} \stackrel{\text{i.i.d.}}{\sim} \Lambda$, to define the random feature map for $\mathbf{x} \in \mathcal{X}$ by

$$\tilde{\varphi}: \mathcal{X} \to \tilde{\mathcal{H}} \coloneqq \mathbb{R}^{2\tilde{d}}, \quad \tilde{\varphi}(\mathbf{x}) \coloneqq \frac{1}{\sqrt{\tilde{d}}} \left(\cos\left(\mathbf{W}^{\mathsf{T}} \mathbf{x}\right), \sin\left(\mathbf{W}^{\mathsf{T}} \mathbf{x}\right) \right), \tag{2}$$

where $\mathbf{W} = (\mathbf{w}_i)_{i=1}^{\tilde{d}} \in \mathbb{R}^{d \times \tilde{d}}$. Then, the corresponding random kernel is defined for $\mathbf{x}, \mathbf{y} \in \mathcal{X}$ as

$$\tilde{\mathbf{k}}: \mathcal{X} \times \mathcal{X} \to \mathbb{R}, \quad \tilde{\mathbf{k}}(\mathbf{x}, \mathbf{y}) = \left\langle \tilde{\varphi}(\mathbf{x}), \tilde{\varphi}(\mathbf{y}) \right\rangle_{\tilde{\mathcal{H}}} = \frac{1}{\tilde{d}} \sum_{i=1}^{\tilde{d}} \cos\left(\mathbf{w}_{i}^{\mathsf{T}}(\mathbf{x} - \mathbf{y})\right)$$
(3)

to provide a probabilistic approximation to k. Indeed, it is a straightforward exercise to check that $k(\mathbf{x}, \mathbf{y}) = \mathbb{E}[\tilde{k}(\mathbf{x}, \mathbf{y})] \approx \tilde{k}(\mathbf{x}, \mathbf{y})$. This approximation converges exponentially fast in \tilde{d} and uniformly over compact subsets of \mathbb{R}^d as proven in [58, Claim 1]. This bound was later tightened and extended to the derivatives of the kernel in the series of works [74, 80, 9], and we provide an adapted version under Theorem B.1 in the supplement.

Tensors and the tensor product. First, we provide a brief overview of tensors and tensor products of vector spaces, which we will use to construct our feature space called the *free algebra over a vector space*. If *U* and *V* are (not necessarily finite-dimensional) vector spaces, their tensor product $U \otimes V$ is a vector space together with a bilinear map $\otimes : U \times V \to U \otimes W$, such that any bilinear map $B : U \times V \to W$ factors through $U \otimes V$, i.e. there exists a unique linear map $\tilde{B} : U \otimes V \to W$, which satisfies $B = \tilde{B} \circ \otimes$. Let \mathcal{B}_U and \mathcal{B}_V respectively be bases of *U* and *V*. Then, $U \times V$ is unique up to isomorphism, and its basis is formed by the tensor product of the bases, i.e. $\mathcal{B}_{U \otimes V} = \{\mathbf{u} \otimes \mathbf{v} \in U \otimes V : \mathbf{u} \in \mathcal{B}_U, \mathbf{v} \in \mathcal{B}_V\}$.

A coordinate-wise construction of the tensor product is given by identifying $U \otimes V$ with the set of functions from $\mathcal{B}_U \times \mathcal{B}_V$ to \mathbb{R} that are nonzero only on finitely many elements, which is a vector space by pointwise operations on functions. Specifically, for $\mathbf{u} \in \mathcal{B}_U$ and $\mathbf{v} \in \mathcal{B}_V$, we may identify $\mathbf{u} \otimes \mathbf{v}$ with the function that takes 1 on (\mathbf{u}, \mathbf{v}) and 0 on every other element of $\mathcal{B}_U \times \mathcal{B}_V$. Then, we can extend this to $U \times V$ by asserting the bilinearity of \otimes , so denoting $\mathbf{x} = \sum_{\mathbf{u} \in \mathcal{B}_U} \alpha_{\mathbf{u}} \mathbf{u}$ and $\mathbf{y} = \sum_{\mathbf{v} \in \mathcal{B}_V} \beta_{\mathbf{v}} \mathbf{v}$, we have

$$\mathbf{x} \otimes \mathbf{y} = \left(\sum_{\mathbf{u} \in \mathcal{B}_U} \alpha_{\mathbf{u}} \mathbf{u}\right) \otimes \left(\sum_{\mathbf{v} \in \mathcal{B}_V} \beta_{\mathbf{v}} \mathbf{v}\right) = \sum_{\mathbf{u} \in \mathcal{B}_U} \sum_{\mathbf{v} \in \mathcal{B}_V} \alpha_{\mathbf{u}} \beta_{\mathbf{v}} \mathbf{u} \otimes \mathbf{v},$$

where finitely many coefficients are nonzero in $\{\alpha_{\mathbf{u}} \in \mathbb{R} : \mathbf{u} \in \mathcal{B}_U\}$ and $\{\beta_{\mathbf{v}} \in \mathbb{R} : \mathbf{v} \in \mathcal{B}_V\}$.

¹A kernel is called translation-invariant if $k(\mathbf{x}, \mathbf{y}) = k(\mathbf{x} + \mathbf{z}, \mathbf{y} + \mathbf{z})$ for any $\mathbf{x}, \mathbf{y}, \mathbf{z} \in \mathbb{R}^d$.

Now, if *U* and *V* are Hilbert spaces, the canonical extension of their inner product to $U \otimes V$ is defined for rank-1 tensors (i.e. tensors of the form $\mathbf{u} \otimes \mathbf{v}$ for $\mathbf{u} \in U$ and $\mathbf{v} \in V$) as

$$\langle \mathbf{u} \otimes \mathbf{v}, \mathbf{x} \otimes \mathbf{y} \rangle_{U \otimes V} \coloneqq \langle \mathbf{u}, \mathbf{x} \rangle_U \langle \mathbf{v}, \mathbf{y} \rangle_V \quad \text{for } \mathbf{u}, \mathbf{x} \in U, \ \mathbf{v}, \mathbf{y} \in V,$$
 (4)

which is then extended to all elements of $U \otimes V$ via bilinearity of the inner product. After completing the space with respect to the induced metric, $U \otimes V$ becomes again a Hilbert space, and we refer the reader to [93, 42] for further details about tensor products.

Free algebras. Now we introduce our feature space \mathcal{H}_{Sig} . That is, we show how to embed a Hilbert space \mathcal{H} into a bigger Hilbert space \mathcal{H}_{Sig} which is also an associative algebra² using a so-called free construction. Since the tensor product is associative, we can unambiguously take take tensor powers of the vector space \mathcal{H} . Denoting $\mathcal{H}^{\otimes m} := \mathcal{H} \otimes \cdots \otimes \mathcal{H}$, we define the free algebra over \mathcal{H} as the set of sequences of tensors indexed by their degree $m \in \mathbb{N}$,

$$\bigoplus_{m\geq 0} \mathcal{H}^{\otimes m} = \left\{ (t_0, \mathbf{t}_1, \mathbf{t}_2, \dots) : \mathbf{t}_m \in \mathcal{H}^{\otimes m} \text{ for } m \in \mathbb{N}, \exists n \in \mathbb{N} \text{ s.t. } N \geq n, \mathbf{t}_N = \mathbf{0} \right\},$$
(5)

where \bigoplus is the direct sum operation, \otimes is the tensor product. For example, if $\mathcal{H} = \mathbb{R}^d$, then the degree-1 component is a *d*-dimensional vector, the degree-2 component is a *d* × *d* matrix, the degree-3 component is an array of shape $d \times d \times d$. The space $\bigoplus_{m\geq 0} \mathcal{H}^{\otimes m}$ is a vector space with addition and scalar multiplication defined for $\lambda \in \mathbb{R}$, **s**, **t** $\in \bigoplus_{m\geq 0} \mathcal{H}^{\otimes m}$ as

$$\mathbf{s} + \mathbf{t} \coloneqq (\mathbf{s}_m + \mathbf{t}_m)_{m \ge 0}, \quad \lambda \mathbf{s} \coloneqq (\lambda \mathbf{s}_m)_{m \ge 0},$$

and \mathcal{H} is a linear subspace of $\bigoplus_{m\geq 0} \mathcal{H}^{\otimes m}$ by identifying $\mathbf{v} \in \mathcal{H}$ as $(0, \mathbf{v}, 0, 0, ...) \in \bigoplus_{m\geq 0} \mathcal{H}^{\otimes m}$. Further, $\bigoplus_{m\geq 0} \mathcal{H}^{\otimes m}$ is also an associative algebra since it is endowed with a (noncommutative³) product defined for tensors $\mathbf{s}, \mathbf{t} \in \bigoplus_{m\geq 0} \mathcal{H}^{\otimes m}$ as

$$\mathbf{st} = \left(\sum_{i=0}^{m} \mathbf{s}_i \otimes \mathbf{t}_{m-i}\right)_{m \ge 0} \in \bigoplus_{m \ge 0} \mathcal{H}^{\otimes m}.$$

This process of turning \mathcal{H} into an algebra $\bigoplus_{m\geq 0} \mathcal{H}^{\otimes m}$ is a free construction; informally this means that (5) is the minimal structure that turns \mathcal{H} into an algebra; for more details about free algebras, see [93, 64]. We now define for $\mathbf{s}, \mathbf{t} \in \bigoplus_{m\geq 0} \mathcal{H}^{\otimes m}$ their inner product as

$$\langle \mathbf{s}, \mathbf{t}
angle_{\oplus_{m \geq 0} \mathcal{H}^{\otimes m}} = \sum_{m \geq 0} \langle \mathbf{s}_m, \mathbf{t}_m
angle_{\mathcal{H}^{\otimes m}}$$
 ,

where the inner product $\langle \mathbf{s}_m, \mathbf{t}_m \rangle_{\mathcal{H}^{\otimes m}}$ on $\mathcal{H}^{\otimes m}$ is as in (4). Finally, the completion of $\bigoplus_{m \ge 0} \mathcal{H}^{\otimes m}$ in this inner product gives a Hilbert space \mathcal{H}_{Sig} , which is equivalently defined as

$$\mathcal{H}_{\text{Sig}} = \{ \mathbf{t} = (t_0, \mathbf{t}_1, \mathbf{t}_2, \ldots) : \mathbf{t}_m \in \mathcal{H}^{\otimes m}, \, \langle \mathbf{t}, \mathbf{t} \rangle_{\mathcal{H}_{\text{Sig}}} < \infty \}.$$
(6)

Path Signatures. A classic way to obtain a structured and hierarchical description of a path $\mathbf{x} : [0, T] \to \mathbb{R}^d$ is by computing a sequence of iterated integrals called the path signature of \mathbf{x} given as tensors of increasing degrees $m \in \mathbb{N}$ such that the degree-*m* object is

$$S_m(\mathbf{x}) \coloneqq \int_{0 < t_1 < \dots < t_m < T} d\mathbf{x}(t_1) \otimes \dots \otimes d\mathbf{x}(t_m) = \int_{0 < t_1 < \dots < t_m < T} \dot{\mathbf{x}}(t_1) \otimes \dots \otimes \dot{\mathbf{x}}(t_m) dt_1 \cdots dt_m.$$

²An algebra *A* is a vector space *A*, where one can multiply elements together, i.e. $\mathbf{ab} \in A$ for $\mathbf{a}, \mathbf{b} \in A$.

³Noncommutative means that $\mathbf{ab} \neq \mathbf{ba}$ in general for elements $\mathbf{a}, \mathbf{b} \in V$ of the algebra.

Formally, we refer to the map that takes a path to its iterated integrals, $S : \mathsf{Paths} \to \mathcal{H}_{Sig}$, $S(\mathbf{x}) \coloneqq (1, S_1(\mathbf{x}), S_2(\mathbf{x}), \ldots)$ as the path signature map. The domain of S is a space of paths that are regular enough such that the integrals are well-defined. Its feature space is given by applying the above construction of \mathcal{H}_{Sig} in (6) to $\mathcal{H} = \mathbb{R}^d$ with the Euclidean inner product.

Among the attractive properties of S is that it linearizes nonlinear functions of paths, that is for any continuous function f one can find a linear functional **w** of S such that

$$f(\mathbf{x}) \approx \langle \mathbf{w}, S(\mathbf{x}) \rangle \coloneqq \sum_{m \ge 0} \sum_{i_1, \dots, i_m \in [d]} w_{i_1, \dots, i_m} \int \dot{\mathbf{x}}^{i_1}(t_1) \cdots \dot{\mathbf{x}}^{i_m}(t_m) \mathrm{d}t_1 \cdots \mathrm{d}t_m, \tag{7}$$

where (7) $w_1, \ldots, w_d, w_{1,1}, \ldots, w_{d,d}, \ldots, w_{d,\ldots,d} \in \mathbb{R}$ denote the coordinates of **w**, and the approximation holds uniformly on compacts [26, Theorem II.5] whenever the path x includes time as a coordinate⁴. The same results generalize to paths without an increasing coordinate up to reparametrization (i.e. time-warping) and backtracking, formally called "treelike" equivalence, see [31]. Moreover, these iterated integrals can be well-defined beyond the setting of smooth paths; for example, the same results extend to Brownian motion, semimartingales, and even rougher paths. Rough path theory provides a systematic study that comes with a rich toolbox, that combines analytic and algebraic estimates, rich enough to cover the trajectories of large classes of stochastic processes; see [51, 27] for an introduction. Informally, iterated integrals of paths can be seen as a generalization of classical monomials and from this perspective, the approximation (7) can be regarded as the extension of classic polynomial regression to path-valued data. Thus at least informally it is not surprising, that vanilla signature features suffer from similar drawbacks as classic monomial features; for example, if classic monomials are replaced by other nonlinearities this often drastically improves the approximations; see e.g. [84, 68], where precomposing the signature with the RBF kernel increases learning performance.

Signature Features for Sequential Data. A challenge in machine learning when constructing feature maps for datasets of sequences is that the sequence length can vary from instance to instance; the space of sequences $\mathcal{X}_{seq} = \{(\mathbf{x}_i)_{i=1}^L : \mathbf{x}_1, \dots, \mathbf{x}_L \in \mathcal{X} \text{ and } L \in \mathbb{Z}_+\}$ includes sequences of various lengths, and they should all get mapped to the same feature space, while preserving the information about the elements themselves and their ordering. A concatenation property of path signatures called Chen's identity [53, Thm. 2.9] turns concatentation into multiplication provides a principled approach to construct features for sequences. Below we recall the construction of discrete-time signatures based on [82].

The key idea is to define the discrete-time signature of 1-step increments, and then glue features together by algebra multiplication to guarantee that the Chen identity holds by construction. Now assume we are given a static feature map $\varphi : \mathfrak{X} \to \mathfrak{H}$ into some Hilbert space \mathfrak{H} . Our task is to construct from this feature map for elements of \mathfrak{X} , a feature map for sequences of arbitrary length in \mathfrak{X} . A natural first step is to apply the feature map φ elementwise to a sequence $\mathbf{x} \in \mathfrak{X}_{seq}$ to lift it to a sequence into the feature space \mathfrak{H} of φ , $\varphi(\mathbf{x}) \coloneqq (\varphi(\mathbf{x}_i))_{i=1}^{\ell_{\mathbf{x}}} \in \mathfrak{H}_{seq}$. The challenge is now to construct a feature map for sequences in \mathfrak{H} . Simple aggregation of the individual features fails; e.g. summation of the individual features $\varphi(\mathbf{x}_i)$ would lose the order information, vectorization $(\varphi(\mathbf{x}_1), \ldots, \varphi(\mathbf{x}_{\ell_x})) \in \mathfrak{H}^{\ell_x}$ would make sequences of different length not comparable. It turns out that multiplication is well-suited for this task in a suitable algebra.

⁴This means that $\mathbf{x}^{i}(t) = t$ for some $i \in [d]$; more generally, a strictly increasing coordinate is sufficient.

Fortunately, there is a natural way to embed any Hilbert space \mathcal{H} into a larger Hilbert space \mathcal{H}_{Sig} that is also a non-commutative algebra. First, we take the 1st-order differences,

$$\mathbf{x} \mapsto \delta \varphi(\mathbf{x}) \coloneqq \left(\varphi(\mathbf{x}_{i+1}) - \varphi(\mathbf{x}_i)\right)_{i=1}^{\ell_{\mathbf{x}} - 1} \in \mathcal{H}^{\ell_{\mathbf{x}} - 1}, \quad \text{where } \mathbf{x} \in \mathcal{X}_{\text{seq}}$$
(8)

since it is more natural to keep track of changes rather than absolute values. Then we identify ${\cal H}$ as a subset of ${\cal H}_{Sig}$. The simplest choice given the above construction of ${\cal H}_{Sig}$ is

$$\iota: \mathbf{h} \mapsto (1, \mathbf{h}, \mathbf{0}, \mathbf{0}, \ldots) \in \mathcal{H}_{\text{Sig}} \quad \text{where } \mathbf{h} \in \mathcal{H}.$$
(9)

A direct calculation shows that composing the maps (8), (9), and multiplying the individual entries in \mathcal{H}_{Sig} results in a sequence summary using all non-contiguous subsequences, since in each multiplication step a sequence entry is either selected once or not at all. This gives rise to the discretized signatures $\varphi_{Sig} : \mathcal{X}_{seq} \to \mathcal{H}_{Sig}$ for $\mathbf{x} \in \mathcal{X}_{seq}$ with $\ell_{\mathbf{x}} \geq 2$:

$$\varphi_{\text{Sig}}(\mathbf{x}) \coloneqq \prod_{i=1}^{\ell_{\mathbf{x}}-1} \iota(\delta\varphi(\mathbf{x}_{i})) = \left(\sum_{\mathbf{i}\in\Delta_{m}(\ell_{\mathbf{x}}-1)} \delta\varphi(\mathbf{x}_{i_{1}}) \otimes \cdots \otimes \delta\varphi(\mathbf{x}_{i_{m}})\right)_{m \ge 0},\tag{10}$$

where $\Delta_m : \mathbb{Z}_+ \to \mathbb{Z}_+^m$ is as defined in (1) and $\mathbf{i} = (i_1, \dots, i_m)$. Thus, the sequence feature is itself a sequence, however, now a sequence of tensors indexed by their degree $m \in \mathbb{N}$ in contrast to being indexed by the time index $i \in [\ell_x]$. These sequence features are invariant to a natural transformation of time series called time-warping, but can also be made sensitive to it by including time as an extra coordinate with the mapping $\mathbf{x} = (\mathbf{x}_i)_{i=1}^{\ell_x} \mapsto (t_i, \mathbf{x}_i)_{i=1}^{\ell_x}$. It also possesses similar approximation properties to path signatures in (7), i.e. uniform approximation of functions of sequences on compact sets; see Appendices A and B in [82].

Despite the abstract derivation, the resulting feature map φ_{Sig} is—in principle—explicitly computable when $\mathcal{H} = \mathbb{R}^d$; see [38] for details. However, when the static feature map φ is high- or infinite-dimensional, this is not feasible and we discuss a kernel trick further below.

Remark 2.1. We used the map ι , as defined in (9), to embed \mathcal{H} into \mathcal{H}_{Sig} . Other choices are possible, for example one could use the embedding $\hat{\iota} : \mathcal{H} \to \mathcal{H}_{Sig}$ for $\mathbf{h} \in \mathcal{H}$

$$\hat{\iota}(\mathbf{h}) \coloneqq \left(1, \mathbf{h}, \frac{\mathbf{h}^{\otimes 2}}{2!}, \frac{\mathbf{h}^{\otimes 3}}{3!}, \ldots\right) \in \mathcal{H}_{\text{Sig}}.$$
(11)

This embedding is actually the classical choice in mathematics, but different choices of the embedding lead to, besides potential improvements in benchmarks, mildly different computational complexities and interesting algebraic questions [20, 82, 83].

Finally, it can be useful to only consider the first $M \in \mathbb{Z}_+$ tensors in the series $\varphi_{\text{Sig}}(\mathbf{x})$ analogously to using the first M moments in classic polynomial regression to avoid overfitting. Hence, we define the M-truncated signature features for $M \in \mathbb{Z}_+$ as

$$\varphi_{\operatorname{Sig}_{\leq M}}(\mathbf{x}) \coloneqq (1, \varphi_{\operatorname{Sig}_{1}}(\mathbf{x}), \dots, \varphi_{\operatorname{Sig}_{M}}(\mathbf{x}), \mathbf{0}, \mathbf{0}, \dots) \quad \text{for } \mathbf{x} \in \mathcal{X}_{\operatorname{seq}},$$

where $\varphi_{\text{Sig}_m}(\mathbf{x})$ is the projection of $\varphi_{\text{Sig}}(\mathbf{x})$ onto $\mathcal{H}^{\otimes m}$. In practice, we regard $M \in \mathbb{Z}_+$, and the choice of the embedding as hyperparameters to optimize.

Signature Kernels. The signature is a powerful feature set for nonlinear regression on paths and sequences. A computational bottleneck associated with it is the dimensionality of the feature space \mathcal{H}_{Sig} . As we are dealing with tensors, for \mathcal{H} finite-dimensional $\varphi_{Sig_m}(\mathbf{x})$ is a tensor of degree-*m* which has $(\dim \mathcal{H})^m$ coordinates that need to be computed. This can quickly become computationally expensive. For infinite-dimensional \mathcal{H} , e.g. when \mathcal{H} is a reproducing kernel Hilbert space (RKHS), which is one of the most interesting settings due to the modelling flexibility, it is infeasible to directly compute φ_{Sig} . In [39], the signature kernel was introduced, and it was shown that a kernel trick allows to compute the inner product of signature features up to a given degree $M \in \mathbb{Z}_+$ using dynamic programming, even when \mathcal{H} is infinite-dimensional. Subsequently, [68] proposed a PDE-based algorithm to compute it, which was further extended in [8], and we refer to [46] for a recent overview of signature kernels. Here, we focus on discrete-time, and our starting point is the approach of [39] combined with the non-geometric approximation [20] resulting in the features (10).

Above we described a generic way to turn a static feature map $\varphi : \mathfrak{X} \to \mathfrak{H}$ into a feature map $\varphi_{\mathrm{Sig}_{\leq M}}(\mathbf{x})$ for sequences, see (10). The signature kernel is a powerful formalism that allows to transform any static kernel on \mathfrak{X} into a kernel for sequences that evolve in \mathfrak{X} . Let $k : \mathfrak{X} \times \mathfrak{X} \to \mathbb{R}$ be a continuous and bounded kernel, and from now on, let \mathfrak{H} denote its RKHS, and $\varphi(\mathbf{x}) \coloneqq \mathsf{k}_{\mathbf{x}} \equiv \mathsf{k}(\mathbf{x}, \cdot)$ the associated reproducing kernel lift for $\mathbf{x} \in \mathfrak{X}$. We define the *M*-truncated (discretized) signature kernel $\mathsf{k}_{\mathrm{Sig}_{\leq M}} : \mathfrak{X}_{\mathrm{seq}} \times \mathfrak{X}_{\mathrm{seq}} \to \mathbb{R}$ for $M \in \mathbb{Z}_+$ as

$$k_{\operatorname{Sig}_{\leq M}}(\mathbf{x}, \mathbf{y}) \coloneqq \left\langle \varphi_{\operatorname{Sig}_{\leq M}}(\mathbf{x}), \varphi_{\operatorname{Sig}_{\leq M}}(\mathbf{y}) \right\rangle_{\mathcal{H}_{\operatorname{Sig}}} = \sum_{m=0}^{M} \left\langle \varphi_{\operatorname{Sig}_{m}}(\mathbf{x}), \varphi_{\operatorname{Sig}_{m}}(\mathbf{y}) \right\rangle_{\mathcal{H}^{\otimes m}}$$
$$= \sum_{m=0}^{M} k_{\operatorname{Sig}_{m}}(\mathbf{x}, \mathbf{y}) = \sum_{m=0}^{M} \sum_{\substack{\mathbf{i} \in \Delta_{m}(\ell_{\mathbf{x}}-1)\\ \mathbf{j} \in \Delta_{m}(\ell_{\mathbf{y}}-1)}} \delta_{i_{1},j_{1}}^{2} k(\mathbf{x}_{i_{1}}, \mathbf{y}_{j_{1}}) \cdots \delta_{i_{m},j_{m}}^{2} k(\mathbf{x}_{i_{m}}, \mathbf{y}_{j_{m}}), \quad (12)$$

where we defined the level-*m* (discretized) signature kernel $k_{\text{Sig}_m} : \mathcal{X}_{\text{seq}} \times \mathcal{X}_{\text{seq}} \to \mathbb{R}$ for $m \in [M]$ as $k_{\text{Sig}_m}(\mathbf{x}, \mathbf{y}) \coloneqq \langle \varphi_{\text{Sig}_m}(\mathbf{x}), \varphi_{\text{Sig}_m}(\mathbf{y}) \rangle_{\mathcal{H}^{\otimes m}}$, and δ^2 denotes a 2nd-order cross-differencing operator such that $\delta_{i,j}^2 k(\mathbf{x}_i, \mathbf{y}_j) \coloneqq k(\mathbf{x}_{i+1}, \mathbf{y}_{j+1}) - k(\mathbf{x}_{i+1}, \mathbf{y}_j) - k(\mathbf{x}_i, \mathbf{y}_{j+1}) + k(\mathbf{x}_i, \mathbf{y}_j)$ for $i \in [\ell_{\mathbf{x}} - 1]$ and $j \in [\ell_{\mathbf{y}} - 1]$. The key insight by [39] is equation (12), i.e. that $k_{\text{Sig}_{\leq M}}$ can be computed⁵ without computing $\varphi_{\text{Sig}_{\leq M}}$ itself by a kernel trick that only uses kernel evaluations.

The kernel hyperparameters are the choice of the static kernel k, for which there is a wide range of options, e.g. for $\mathcal{X} = \mathbb{R}^d$ the Gaussian, exponential or Matérn family of kernels; any hyperparameters that k comes with, such as the bandwidth; the truncation level $M \in \mathbb{Z}_+$; the choice of the algebra embedding, e.g. (9) or (11); and the choice of kernel normalization [10] that scales each level k_{Sig_m} appropriately. It also comes with nice theoretical guarantees such as analytic estimates when sequences converge to paths, its maximum mean discrepancy (MMD) metrizes classic topologies for stochastic processes, and can lead to robust statistics in the classic statistical sense (B-robustness); see [10] for details.

Although (12) looks expensive to compute, [39] applies dynamic programming to efficiently compute $k_{Sig_{\leq M}}$ using a recursive algorithm; an alternative algorithm is the above mentioned approach of approximating the (untruncated) signature kernel k_{Sig} using PDE-discretization. Importantly, (12) avoids computing tensors, and only depends on the entrywise evaluations of the static kernel $k(\mathbf{x}_i, \mathbf{y}_j)$. Indeed, this leads to a computational cost of $O((M + d)\ell_{\mathbf{x}}\ell_{\mathbf{y}})$, which is feasible for sequences evolving in high-dimensional state-spaces, but only with moderate sequence length. Note that the same bottleneck applies to PDE-based approaches. In part, the aim of this article is to alleviate this quadratic cost in se-

⁵The computation can be carried out exactly for finite *M* and approximately for $M = \infty$.

quence length, while approximately enjoying the modelling capability of working within an infinite-dimensional RKHS.

3 Random Fourier Signature Features

The goal of this section is to build random features for sequences, that enjoy the benefit of linear sequence length and low-dimensional feature complexity with theoretical guarantees that the corresponding inner product is close to the *M*-truncated (discretized) signature kernel $k_{Sig_{\leq M}}$ with high probability. We construct these random features in a two step process: firstly, we reduce the feature space from infinite to finite (but high) dimensionality through a careful construction using random Fourier features (RFFs), and in the second step we apply further dimensionality reduction to reduce the complexity to an even lower dimensional space in order to aid in scalability. Although we present this construction as conceptually distinct steps, the steps are coupled during the computation, and the features can be computed directly without going through the initial step.

From infinite to finite dimensions. In Section 2, we recalled the RFF construction, which associates to a continuous, bounded, translation-invariant kernel $k : \mathcal{X} \times \mathcal{X} \to \mathbb{R}$ on \mathcal{X} a spectral measure Λ , and approximates k by drawing samples from Λ to define the random features $\tilde{\varphi} : \mathcal{X} \to \tilde{\mathcal{H}}$ (2), and the random kernel $\tilde{k} : \mathcal{X} \times \mathcal{X} \to \mathbb{R}$ (3). Afterwards, we presented a generic way to turn any such static features $\tilde{\varphi} : \mathcal{X} \to \tilde{\mathcal{H}}$ for elements of \mathcal{X} into sequence features for sequences that evolve in \mathcal{X} via $\varphi_{\mathrm{Sig}_{\leq M}} : \mathcal{X}_{\mathrm{seq}} \to \mathcal{H}_{\mathrm{Sig}}$. Applying this construction with the RFF as feature map on \mathcal{X} would already result in a random feature map for sequences, i.e. a map from $\mathcal{X}_{\mathrm{seq}}$ into $\tilde{\mathcal{H}}_{\mathrm{Sig}}$. Taking the inner product in $\tilde{\mathcal{H}}_{\mathrm{Sig}}$ of this new random feature map for sequences would, however, only yield a biased estimator for the truncated signature kernel $k_{\mathrm{Sig}_{\leq M}}$. We correct for this bias by revisiting our previous construction, and build an unbiased approximation to $k_{\mathrm{Sig}_{\leq M}}$ using independent RFF copies in each tensor multiplication step. Then, we show in Theorem 3.2 that this random estimator comes with good probabilistic guarantees.

The probabilistic construction procedure is outlined in the following definition.

Definition 3.1. Let $\mathbf{W}^{(1)}, \dots, \mathbf{W}^{(M)} \stackrel{i.i.d.}{\sim} \Lambda^{\tilde{d}}$ be i.i.d. random matrices sampled from $\Lambda^{\tilde{d}}$ for RFF dimension $\tilde{d} \in \mathbb{Z}_+$, and the define the independent RFF maps $\tilde{\varphi}_m : \mathfrak{X} \to \tilde{\mathfrak{H}}$ as in (2), i.e. $\tilde{\varphi}_m(\mathbf{x}) = \frac{1}{\sqrt{\tilde{d}}} \left(\cos(\mathbf{W}^{(m)^{\top}} \mathbf{x}), \sin(\mathbf{W}^{(m)^{\top}} \mathbf{x}) \right)$ for $m \in [M]$ and $\mathbf{x} \in \mathfrak{X}$. The *M*-truncated Random Fourier Signature Feature (RFSF) map $\tilde{\varphi}_{\operatorname{Sig}_{\leq M}} : \mathfrak{X}_{\operatorname{seq}} \to \tilde{\mathfrak{H}}_{\operatorname{Sig}}$ from sequences in \mathfrak{X} into the free algebra over $\tilde{\mathfrak{H}}$ is defined for truncation level $M \in \mathbb{Z}_+$ and $\mathbf{x} \in \mathfrak{X}_{\operatorname{seq}}$ as

$$\tilde{\varphi}_{\text{Sig}_{\leq M}}(\mathbf{x}) \coloneqq \left(\sum_{\mathbf{i} \in \Delta_m(\ell_{\mathbf{x}} - 1)} \delta \tilde{\varphi}_1(\mathbf{x}_{i_1}) \otimes \cdots \otimes \delta \tilde{\varphi}_m(\mathbf{x}_{i_m}) \right)_{m=0}^M.$$
(13)

Further, the RFSF *kernel* $\tilde{k}_{\text{Sig}_{\leq MM}}$: $\mathfrak{X}_{\text{seq}} \times \mathfrak{X}_{\text{seq}} \rightarrow \mathbb{R}$ *can be computed for* $\mathbf{x}, \mathbf{y} \in \mathfrak{X}_{\text{seq}}$ *as*

$$\widetilde{\mathsf{k}}_{\mathrm{Sig}_{\leq M}}(\mathbf{x}, \mathbf{y}) \coloneqq \left\langle \widetilde{\varphi}_{\mathrm{Sig}_{\leq M}}(\mathbf{x}), \widetilde{\varphi}_{\mathrm{Sig}_{\leq M}}(\mathbf{y}) \right\rangle_{\widetilde{\mathcal{H}}_{\mathrm{Sig}}} = \sum_{m=0}^{M} \left\langle \widetilde{\varphi}_{\mathrm{Sig}_{m}}(\mathbf{x}), \widetilde{\varphi}_{\mathrm{Sig}_{m}}(\mathbf{y}) \right\rangle_{\widetilde{\mathcal{H}}^{\otimes m}} \\
= \sum_{m=0}^{M} \widetilde{\mathsf{k}}_{\mathrm{Sig}_{m}}(\mathbf{x}, \mathbf{y}) = \sum_{m=0}^{M} \sum_{\substack{\mathbf{i} \in \Delta_{m}(\ell_{\mathbf{x}}-1)\\ \mathbf{j} \in \Delta_{m}(\ell_{\mathbf{y}}-1)}} \delta_{i_{1}, j_{1}}^{2} \widetilde{\mathsf{k}}_{1}(\mathbf{x}_{i_{1}}, \mathbf{y}_{j_{1}}) \cdots \delta_{i_{m}, j_{m}}^{2} \widetilde{\mathsf{k}}_{m}(\mathbf{x}_{i_{m}}, \mathbf{y}_{j_{m}}), \quad (14)$$

where we defined the level-m RFSF kernel $\tilde{k}_{\text{Sig}_m} : \mathfrak{X}_{\text{seq}} \times \mathfrak{X}_{\text{seq}} \to \mathbb{R}$ for $m \in \mathbb{N}$ as $\tilde{k}_{\text{Sig}_m}(\mathbf{x}, \mathbf{y}) \coloneqq \langle \tilde{\varphi}_{\text{Sig}_m}(\mathbf{x}), \tilde{\varphi}_{\text{Sig}_m}(\mathbf{y}) \rangle_{\tilde{\mathcal{H}}^{\otimes m}}$ with the convention that $\tilde{k}_{\text{Sig}_0} \equiv 1$, and $\tilde{k}_1, \dots, \tilde{k}_M : \mathfrak{X} \times \mathfrak{X} \to \mathbb{R}$ are independent RFF kernels defined as in (3) with the random weights $\mathbf{W}^{(1)}, \dots, \mathbf{W}^{(M)} \in \mathbb{R}^{d \times \tilde{d}}$.

Since the feature map $\tilde{\varphi}_{\text{Sig}_{\leq M}}$ can be directly evaluated in the feature space recursively, $\tilde{k}_{\text{Sig}_{\leq M}}$ has linear complexity in the sequence length. However, it requires computing highdimensional tensors, where the degree-*m* component $\tilde{\varphi}_{\text{Sig}_m}(\mathbf{x}) \in \tilde{\mathcal{H}}^{\otimes m}$ has $(\dim \tilde{\mathcal{H}})^m = (2\tilde{d})^m$ coordinates, making it infeasible for large $m, \tilde{d} \in \mathbb{Z}_+$. Remark 3.3 discusses the computational complexity in detail. Further, note that the kernel can be evaluated by means of a kernel trick exactly analogously to the evaluation of (12), but in this case there are no computational gains compared to the infinite-dimensional signature kernel $k_{\text{Sig}_{\leq M}}(\mathbf{x}, \mathbf{y})$.

Next, we provide a theoretical analysis to show that the random kernel $\tilde{k}_{Sig_{\leq M}}(\mathbf{x}, \mathbf{y})$ converges to the ground truth signature kernel $k_{Sig_{\leq M}}(\mathbf{x}, \mathbf{y})$ exponentially fast and uniformly over compact state-spaces $\mathfrak{X} \subseteq \mathbb{R}^d$, generalizing the result [58, Claim 2] to this non-Euclidean domain of sequences. Throughout the analysis, we need certain regularity properties of Λ in order to invoke quantitative versions of the law of large numbers, i.e. properties such as boundedness, existence of the moment-generating function, moment-boundedness, or belonging to certain Orlicz spaces of random variables. Boundedness of the spectral measure is too restrictive an assumption, since a continuous, bounded, translation-invariant kernel $k : \mathfrak{X} \times \mathfrak{X} \to \mathbb{R}$ is characteristic if and only if the support of its spectral measure is \mathbb{R}^d , see [73, Prop. 8]. Hence, we instead work with the assumption that its moments are well-controllable, i.e. the tails of the distribution are not "too heavy". Specifically, we assume the Bernstein moment condition that

$$\mathbb{E}_{\mathbf{w} \sim \Lambda} \left[w_i^{2m} \right] \le \frac{m! S^2 R^{m-2}}{2} \quad \text{for all } i \in [d]$$
(15)

for some *S*, *R* > 0. We show in the Supplementary Material under Lemmas A.11 and A.12, in a more general context, that this is equivalent to Λ being a sub-Gaussian probability measure; see e.g. [4, Sec 2.3] and [87, Sec. 2.5] about sub-Gaussianity. This of course includes the spectral measure of the Gaussian kernel defined for bandwidth $\sigma > 0$ and $\mathbf{x}, \mathbf{y} \in \mathcal{X}$ k(\mathbf{x}, \mathbf{y}) = exp $\left(-\|\mathbf{x}-\mathbf{y}\|_{2}^{2}/2\sigma^{2}\right)$, which has a Gaussian spectral distribution $\mathbf{w} \sim \mathcal{N}(0, 1/\sigma^{2}\mathbf{I}_{d})$, and therefore calculation gives $\mathbb{E}_{w \sim \mathcal{N}(0, 1/\sigma^{2})}\left[w^{2m}\right] = \frac{2^{m}\Gamma(m+\frac{1}{2})}{\sigma^{2}m\sqrt{\pi}} < \frac{m!}{2}\left(\frac{2\sqrt{2}}{\sigma^{2}}\right)^{m-2}$, since $\Gamma(m+1/2) < \Gamma(m+1) = m!$. Hence Λ satisfies condition (15) with *S*, *R* as given here.

Now we state our approximation theorem regarding \tilde{k}_{Sig_m} , which quantifies that it is a (sub-)exponentially good estimator of k_{Sig_m} with high probability and uniformly.

Theorem 3.2. Let $k : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ be a continuous, bounded, translation-invariant kernel with spectral measure Λ , which satisfies (15). Let $\mathfrak{X} \subset \mathbb{R}^d$ be compact and convex with diameter $|\mathfrak{X}|, \mathfrak{X}_\Delta \coloneqq \{\mathbf{x} - \mathbf{y} : \mathbf{x}, \mathbf{y} \in \mathfrak{X}\}$. Then, the following quantities are finite: $\sigma_{\Lambda}^2 \coloneqq \mathbb{E}_{\mathbf{w} \sim \Lambda}[\|\mathbf{w}\|_2^2]$, $L \coloneqq \|\mathbb{E}_{\mathbf{w} \sim \Lambda}[\mathbf{w}\mathbf{w}^{\top}]\|_2^{1/2}$, $E_{i,j} \coloneqq \mathbb{E}_{\mathbf{w} \sim \Lambda}[\|\mathbf{w}\|_2]$ and $D_{i,j} \coloneqq \sup_{\mathbf{z} \in \mathfrak{X}_\Delta} \|\nabla \begin{bmatrix} \frac{\partial^2 k(\mathbf{z})}{\partial z_i \partial z_j} \end{bmatrix}\|_2$ for $i, j \in [d]$. Further, for any max. sequence 1-var V > 0, and signature level $m \in \mathbb{Z}_+$, for $\epsilon > 0$

$$\mathbb{P}\left|\sup_{\substack{\mathbf{x},\mathbf{y}\in\mathcal{X}_{\text{seq}}\\||\mathbf{x}||_{1-\text{var}},||\mathbf{y}||_{1-\text{var}}\leq V}}\left|\mathsf{k}_{\text{Sig}_{m}}(\mathbf{x},\mathbf{y})-\tilde{\mathsf{k}}_{\text{Sig}_{m}}(\mathbf{x},\mathbf{y})\right|\geq\epsilon\right|\leq \left|\left|\left|\mathsf{x}\right|\right|_{1-\text{var}}\leq V\right|\right|^{\frac{1}{d+1}}\leq V \\ \leq m\left\{\left(C_{d,\mathcal{X}}\left(\frac{\beta_{d,m,V}}{\epsilon}\right)^{\frac{d}{d+1}}+d\right)\exp\left(-\frac{\tilde{d}}{2(d+1)(S^{2}+R)}\left(\frac{\epsilon}{\beta_{d,m,V}}\right)^{2}\right) \quad \text{for } \epsilon<\beta_{d,m,V} \\ \left(C_{d,\mathcal{X}}\left(\frac{\beta_{d,m,V}}{\epsilon}\right)^{\frac{d}{(d+1)m}}+d\right)\exp\left(-\frac{\tilde{d}}{2(d+1)(S^{2}+R)}\left(\frac{\epsilon}{\beta_{d,m,V}}\right)^{\frac{1}{m}}\right) \quad \text{for } \epsilon\geq\beta_{d,m,V}, \end{aligned}\right. \tag{16}$$

where $C_{d,\mathfrak{X}} \coloneqq 2^{\frac{1}{d+1}} 16 |\mathfrak{X}|^{\frac{d}{d+1}} \sum_{i,j=1}^{d} (D_{i,j} + E_{i,j})^{\frac{d}{d+1}} and \beta_{d,m,V} \coloneqq m \left(2V^2 \left(L^2 \vee 1\right) \left(\sigma_{\Lambda}^2 \vee d\right)\right)^m$.

The proof is provided in the supplement under Theorem C.10. The result shows that the random kernel \tilde{k}_{Sig_m} approximates the signature kernel k_{Sig_m} uniformly over subsets of \mathcal{X}_{seq} of sequences $\mathbf{x} \in \mathcal{X}_{seq}$ with maximal 1-variation V, $\|\mathbf{x}\|_{1-var} \leq V$, assuming that the statespace $\mathcal{X} \subset \mathbb{R}^d$ is a convex and compact domain. The error bound is analogous to the classic RFF bounds, in the sense that the tail probability decreases exponentially fast as a function of the RFF dimension \tilde{d} . The functional form of the bound is inherited from Theorem B.1, which provides an analogous result for the derivatives of RFF. This link follows from Lemma C.9, which connects the concentration of the RFSF kernel to the second derivatives of RFF.

The main difference from the classic case, i.e. [58, Claim 1] and Theorem B.1, is the appearance of $\beta_{d,m,V}$ which controls a regime change in the tail behaviour. Concretely, for $\epsilon < \beta_{d,m,V}$ (16) has a polynomial plus a sub-Gaussian tail, while for $\epsilon > \beta_{d,m,V}$ has a (1/*m*)-subexponential tail. This is not surprising as the inner summand in (14) is the *m*-fold tensor product of *m* independent RFF kernels, which makes the tail heavier exactly by an exponent of 1/m. The constant itself, $\beta_{d,m,V}$, depends on (i) the maximal sequence 1-variation *V*, which measures a notion of time-warping invariant sequence complexity; (ii) the Lipschitz constant of the kernel *L* (see Examples C.2 and C.3); (iii) the trace of the second moment of Λ , $\sigma_{\Lambda}^2 = \mathbb{E}_{\mathbf{w} \sim \Lambda} [||\mathbf{w}||_2^2]$; (iv) the state-space dimension *d*; (v) and the signature level *m* itself.

Remark 3.3. Algorithm 1 demonstrates the computation of the RFSF map $\tilde{\varphi}_{Sig_{\leq M}}$ given a dataset of sequences $\mathbf{X} = (\mathbf{x}_i)_{i=1}^N \subset \mathfrak{X}_{seq}$. Upon inspection, we can deduce that the algorithm has a computational complexity of $O(NL(Md\tilde{d}+1+\tilde{d}+\ldots+\tilde{d}^M))$. Importantly, it is linear in L, the sequence length, although scales polynomially in the RFF sample size \tilde{d}^M .

Dimensionality Reduction: Diagonal Projection. Previously, we introduced a featurized approximation $\tilde{k}_{Sig_{\leq M}}$ to the signature kernel $k_{Sig_{\leq M}}$, called the RFSF kernel, which reduces the computation from the infinite-dimensional RKHS to a finite-dimensional feature space using random tensors. Although this makes the computation in the feature space viable of the RFSF map $\tilde{\varphi}_{Sig_{\leq M}}$, it is still tensor-valued, which incurs a computational cost of $O(\tilde{d} + \tilde{d}^2 + \cdots + \tilde{d}^m)$ in the RFF dimension $\tilde{d} \in \mathbb{Z}_+$. Now, we take another step towards scalability and apply further dimensionality reduction. By examining the structure of these tensors, we introduce a diagonally projected variant called RFSF-DP that considerably reduces their sizes. We emphasize that the above RFSF construction is the crucial step: it approximates the inner product in an infinite-dimensional space, and now we further approximate it in an even lower dimensional space. The benefit is that one does not have to go through the computation of the initial RFSF map, but only the selected degrees of freedom have to be computed from the beginning.

As a first observation, we notice that the computation of (14) can be reformulated, due to (3) and linearity of the differencing operator, in the following way:

$$\tilde{\mathsf{k}}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y}) = \frac{1}{\tilde{d}^{m}} \sum_{\substack{q_{1},\dots,q_{m}=1\\ \mathbf{j}\in\Delta_{m}(\ell_{\mathbf{y}}-1)}}^{\tilde{d}} \sum_{\substack{p=1\\ p\in\Delta_{m}(\ell_{\mathbf{y}}-1)}}^{m} \delta_{i_{p},j_{p}}^{2} \cos\left(\mathbf{w}_{q_{p}}^{(p)^{\mathsf{T}}}(\mathbf{x}_{i_{p}}-\mathbf{y}_{j_{p}})\right)$$
(17)

by spelling out the definition of the RFF kernel, where $\mathbf{w}_1^{(1)}, \ldots, \mathbf{w}_{\tilde{d}}^{(m)} \stackrel{\text{i.i.d.}}{\sim} \Lambda$, such that $\mathbf{W}^{(p)} = \left(\mathbf{w}_1^{(p)}, \ldots, \mathbf{w}_{\tilde{d}}^{(p)}\right) \in \mathbb{R}^{d \times \tilde{d}}$ as defined in Def. 3.1. Now, we may observe that there that there is a dependency structure among the samples being averaged in (17), since the outer summation is over the Cartesian product $(q_1, \ldots, q_m) \in [\tilde{d}]^{\times m}$, which suggests that we might be

able to drastically reduce the degrees of freedom by restricting this summation to only go over an independent set of samples. One way to do this is to restrict to multi-indices of the form $\mathcal{I} := \{(q, \ldots, q) \in [\tilde{d}]^{\times m} : q \in [\tilde{d}]\}$, i.e. we diagonally project the index set, motivating the name of the approach stated in the following definition.

Definition 3.4. Let $\mathbf{w}_1^{(1)}, \dots, \mathbf{w}_{\tilde{d}}^{(M) \ i.i.d.} \wedge \text{ for } \tilde{d} \in \mathbb{Z}_+, \text{ and define } \hat{\varphi}_{m,q} : \mathfrak{X} \to \hat{\mathfrak{H}} \coloneqq \mathbb{R}^2 \text{ with sample size } \hat{d} = 1 \text{ for } q \in [\tilde{d}] \text{ and } m \in [M], \text{ such that } \hat{\varphi}_{m,q}(\mathbf{x}) = \left(\cos(\mathbf{w}_q^{(m)^{\top}}\mathbf{x}), \sin(\mathbf{w}_q^{(m)^{\top}}\mathbf{x})\right) \text{ for } \mathbf{x} \in \mathfrak{X}.$ The *M*-truncated Diagonally Projected Random Fourier Signature Feature (RFSF-DP) map $\tilde{\varphi}_{\operatorname{Sig}_{\leq M}}^{\operatorname{DP}} : \mathfrak{X}_{\operatorname{seq}} \to \tilde{\mathfrak{H}}_{\operatorname{Sig}}^{\operatorname{DP}} \coloneqq \bigoplus_{m=0}^{M} (\hat{\mathfrak{H}}^{\otimes m})^{\tilde{d}} \text{ is defined for truncation } M \in \mathbb{Z}_+ \text{ and } \mathbf{x} \in \mathfrak{X}_{\operatorname{seq}} \text{ as}$

$$\tilde{\varphi}_{\mathrm{Sig}_{\leq M}}^{\mathrm{DP}}(\mathbf{x}) \coloneqq \frac{1}{\sqrt{\tilde{d}}} \left(\left(\sum_{\mathbf{i} \in \Delta_m(\ell_{\mathbf{x}}-1)} \delta \, \hat{\varphi}_{1,q}(\mathbf{x}_{i_1}) \otimes \cdots \otimes \delta \, \hat{\varphi}_{m,q}(\mathbf{x}_{i_m}) \right)_{q=1}^{\tilde{d}} \right)_{m=0}^{M}.$$

Then, the RFSF-DP kernel can be directly computed for $\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq}$ via

$$\tilde{\mathbf{k}}_{\mathrm{Sig}_{\leq M}}^{\mathrm{DP}}(\mathbf{x}, \mathbf{y}) \coloneqq \left\langle \tilde{\varphi}_{\mathrm{Sig}_{\leq M}}^{\mathrm{DP}}(\mathbf{x}), \tilde{\varphi}_{\mathrm{Sig}_{\leq M}}^{\mathrm{DP}}(\mathbf{y}) \right\rangle_{\tilde{\mathcal{H}}_{\mathrm{Sig}}^{\mathrm{DP}}} = \sum_{m=0}^{M} \left\langle \tilde{\varphi}_{\mathrm{Sig}_{m}}^{\mathrm{DP}}(\mathbf{x}), \tilde{\varphi}_{\mathrm{Sig}_{m}}^{\mathrm{DP}}(\mathbf{y}) \right\rangle_{\left(\hat{\mathcal{H}}^{\otimes m}\right)^{\tilde{d}}} = \sum_{m=0}^{M} \tilde{\mathbf{k}}_{\mathrm{Sig}_{m}}^{\mathrm{DP}}(\mathbf{x}, \mathbf{y})$$

$$= \frac{1}{\tilde{d}} \sum_{m=0}^{M} \sum_{q=1}^{\tilde{d}} \sum_{\substack{\mathbf{i} \in \Delta_{m}(\ell_{\mathbf{x}}-1)\\ \mathbf{j} \in \Delta_{m}(\ell_{\mathbf{y}}-1)}} \delta_{i_{1},j_{1}}^{2} \hat{\mathbf{k}}_{1,q}(\mathbf{x}_{i_{1}}, \mathbf{y}_{j_{1}}) \cdots \delta_{i_{m},j_{m}}^{2} \hat{\mathbf{k}}_{m,q}(\mathbf{x}_{i_{m}}, \mathbf{y}_{j_{m}}),$$
(18)

where we defined the level-m RFSF-DP kernel $\tilde{k}_{\text{Sig}_m}^{\text{DP}} : \mathfrak{X}_{\text{seq}} \times \mathfrak{X}_{\text{seq}} \to \mathbb{R}$ for $m \in \mathbb{N}$ and $\mathbf{x}, \mathbf{y} \in \mathfrak{X}_{\text{seq}}$ as $\tilde{k}_{\text{Sig}}(\mathbf{x}, \mathbf{y}) \coloneqq \left\langle \tilde{\varphi}_{\text{Sig}_m}^{\text{DP}}(\mathbf{x}), \tilde{\varphi}_{\text{Sig}_m}^{\text{DP}}(\mathbf{y}) \right\rangle_{(\hat{\mathcal{H}}(\otimes m))^{\tilde{d}}}$ with the convention that $\tilde{k}_{\text{Sig}_0}^{\text{DP}} \equiv 1$, and $\hat{k}_{m,q} : \mathfrak{X} \times \mathfrak{X} \to \mathbb{R}$ are independent RFF kernels with sample size $\hat{d} = 1$ defined for $\mathbf{x}, \mathbf{y} \in \mathfrak{X}$ as $\hat{k}_{m,q}(\mathbf{x}, \mathbf{y}) \coloneqq \left\langle \hat{\varphi}_{m,q}(\mathbf{x}), \hat{\varphi}_{m,q}(\mathbf{y}) \right\rangle_{\hat{\mathcal{H}}}$ with the random weights $\mathbf{w}_q^{(m)} \in \mathbb{R}^d$ for $q \in [\tilde{d}], m \in [M]$.

Note that by the definition of the RFF kernels in (18), we may substitute that $\hat{k}_{p,q}(\mathbf{x}, \mathbf{y}) = \cos(\mathbf{w}_q^{(p)^{\top}}(\mathbf{x}-\mathbf{y}))$ for $\mathbf{x}, \mathbf{y} \in \mathcal{X}$, so (18) is equivalently written for $\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq}$ as

$$\tilde{\mathsf{K}}_{\text{Sig}_m}^{\text{DP}}(\mathbf{x}, \mathbf{y}) = \frac{1}{\tilde{d}} \sum_{\substack{q=1 \ \mathbf{j} \in \Delta_m(\ell_{\mathbf{x}}-1) \\ \mathbf{j} \in \Delta_m(\ell_{\mathbf{y}}-1)}} \delta_{i_1, j_1}^2 \cos(\mathbf{w}_q^{(1)^{\mathsf{T}}}(\mathbf{x}_{i_1} - \mathbf{y}_{j_1})) \cdots \delta_{i_m, j_m}^2 \cos(\mathbf{w}_q^{(m)^{\mathsf{T}}}(\mathbf{x}_{i_m} - \mathbf{y}_{j_m})),$$

which is what we set out to do in the above paragraph; that is, restrict the outer summation onto the diagonal projection of the index set. Another way to look at Definition 3.4 is that the RFSF-DP kernel in (18) is constructed by defining \tilde{d} independent RFSF kernels, each with internal RFF sample size $\hat{d} = 1$, and then taking their average; the concatenation of their corresponding features are then the features of the RFSF-DP map. Note that for RFF sample size 1, each RFF map has dimension 2, i.e. $\hat{\mathcal{H}} = \mathbb{R}^2$, and hence, the corresponding RFSF kernels have dimension $1+2+\dots+2^M = (2^{M+1}-1)$, which by concatenation results in the overall dimensionality of the RFSF-DP kernel being dim $\tilde{\mathcal{H}}_{Sig}^{TRP} = \tilde{d}(2^{M+1}-1)$. This relates to the computational complexity of the RFSF-DP map; for details see Remark 3.6.

Next, we state our concentration result regarding the level-*m* RFSF-DP kernel $\tilde{k}_{Sig...}^{DP}(\mathbf{x}, \mathbf{y})$.

Theorem 3.5. Let $k : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ be a continuous, bounded, translation-invariant kernel with spectral measure Λ , which satisfies (15). Then, for level $m \in \mathbb{Z}_+$, $\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq}$, and $\epsilon > 0$

$$\mathbb{P}\left[\left|\tilde{\mathsf{K}}_{\operatorname{Sig}_{m}}^{\mathsf{DP}}(\mathbf{x},\mathbf{y})-\mathsf{k}_{\operatorname{Sig}_{m}}(\mathbf{x},\mathbf{y})\right|\geq\epsilon\right]\leq2\exp\left(-\frac{1}{4}\min\left\{\left(\frac{\sqrt{\tilde{d}}\epsilon}{2C_{d,m,\mathbf{x},\mathbf{y}}}\right)^{2},\left(\frac{\tilde{d}\epsilon}{\sqrt{8}C_{d,m,\mathbf{x},\mathbf{y}}}\right)^{\frac{1}{m}}\right\}\right),$$

where $L := \|\mathbb{E}_{\mathbf{w} \sim \Lambda} [\mathbf{w} \mathbf{w}^{\top}]\|$ is the Lipschitz constant of k, and $C_{d,m,\mathbf{x},\mathbf{y}} > 0$ is bounded by

$$C_{d,m,\mathbf{x},\mathbf{y}} \leq \sqrt{8}e^{4}(2\pi)^{1/4}e^{1/24}(4e^{3}\|\mathbf{x}\|_{1-\operatorname{var}}\|\mathbf{y}\|_{1-\operatorname{var}}/m)^{m}\left((2d\max(S,R))^{m} + (L^{2}/\ln 2)^{m}\right).$$

The proof is provided in the supplement under Theorem C.11. The result shows that the RFSF-DP kernel converges for any two sequences $\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq}$ with a $\left(\frac{1}{m}\right)$ -subexponential convergence rate with respect to the sample size $\tilde{d} \in \mathbb{Z}_+$. Similarly to Theorem 3.2, the bound has a phase transition, where for small values of ϵ , it has a sub-Gaussian tail, while for larger values, it has a $\left(\frac{1}{m}\right)$ -subexponential tail. A crucial difference from the previous bound is that now the phase transition happens at $\epsilon^* = C_{d,m,\mathbf{x},\mathbf{y}} 2^{\frac{2m-3/2}{2m-1}} \tilde{d}^{\frac{1-m}{2m-1}}$, which depends on the sample size \tilde{d} . This means that for fixed value of $\epsilon > 0$, the phase transition always happens eventually as \tilde{d} gets large enough, hence the convergence rate with respect to \tilde{d} is $\left(\frac{1}{m}\right)$ -subexponential regardless of the value of ϵ . The slightly reduced rate of convergence compared to the RFSF kernel in Theorem 3.2 is to be expected, since the sample size of the RFSF-DP kernel is analogously reduced by an exponent of $\left(\frac{1}{m}\right)$ in comparison. The constant $C_{d,m,\mathbf{x},\mathbf{y}}$, similarly to (16), depends on (i) the 1-variation of sequences $\|\mathbf{x}\|_{1-\text{var}}$, $\|\mathbf{y}\|_{1-\text{var}}$ that measure the complexity of the sequences; (ii) L > 0, the Lipschitz constant of the kernel k (see Examples C.2, C.3); (iii) the moment bound parameters S, R > 0 from condition (15); (iv) the state-space dimension d; and (v) the signature level m.

Remark 3.6. Algorithm 2 demonstrates the computation of the RFSF-DP map $\tilde{\varphi}_{\text{Sig}_{\leq M}}^{\text{DP}}$ given a dataset of sequences $\mathbf{X} = (\mathbf{x}_i)_{i=1}^N \subset \mathcal{X}_{\text{seq}}$. Upon counting the operations, we deduce that the algorithm has a computational complexity $O(NL\tilde{d}(Md+2^M))$. Crucially, it is linear in both L, the maximal sequence length, and \tilde{d} , the sample size of the random kernel.

Dimensionality Reduction: Tensor Random Projection. Previously, we built the RFSF-DP map by subsampling an independent set from the samples that constitute RFSF kernel. Here, we propose an alternative dimensionality reduction technique that starts again from the RFSF map, and uses random projections to project this generally high-dimensional tensor onto a lower dimension. Random projections are a classic technique in data science for reducing the data dimension, while preserving its important structural properties. They are built upon the celebrated Johnson-Lindenstrauss lemma [36], which states that a set of points in a high-dimensional space can be embedded into a space of much lower dimension, while approximately preserving their geometry. Exploiting this property, we construct a tensor random projected (TRP) variant of our random kernel called RFSF-TRP, such that the computation is coupled between the RFSF and TRP maps, similarly to a kernel trick.

Tensorized random projections [78, 61] construct random projections for tensors with concise parametrization that respects their tensorial nature. Given tensors $\mathbf{s}, \mathbf{t} \in (\mathbb{R}^d)^{\otimes m}$ for $m \in \mathbb{Z}_+$, the TRP map with CP (CANDECOMP/PARAFAC [40]) rank-1 is built via a random functional $\Pr : (\mathbb{R}^d)^{\otimes m} \to \mathbb{R}$ such that $\Pr(\mathbf{s}) = \langle \mathbf{p}_1 \otimes \cdots \otimes \mathbf{p}_m, \mathbf{s} \rangle_{(\mathbb{R}^d)^{\otimes m}}$, where $\mathbf{p}_1, \dots, \mathbf{p}_m \overset{\text{i.i.d.}}{\sim} \mathcal{N}(0, \mathbf{I}_d)$ are *d*-dimensional component vectors sampled from a standard normal distribution. Then, the inner product can be estimated as $\Pr(\mathbf{s})\Pr(\mathbf{t}) \approx \mathbb{E}[\Pr(\mathbf{s})\Pr(\mathbf{t})] = \langle \mathbf{s}, \mathbf{t} \rangle_{(\mathbb{R}^d)^{\otimes m}}$.

Variance reduction is achieved by stacking $n \in \mathbb{Z}_+$ such random projections, each with i.i.d. component vectors $\mathbf{p}_1^{(1)}, \dots, \mathbf{p}_m^{(n) \text{ i.i.d.}} \mathcal{N}(0, \mathbf{I}_d)$. Hence, the TRP operator is defined as

$$\mathsf{TRP}: \left(\mathbb{R}^{d}\right)^{\otimes m} \to \mathbb{R}^{n}, \quad \mathsf{TRP}(\mathbf{s}) \coloneqq \frac{1}{\sqrt{n}} \left(\left\langle \mathbf{p}_{1}^{(i)} \otimes \cdots \otimes \mathbf{p}_{m}^{(i)}, \mathbf{s} \right\rangle \right)_{i=1}^{n}.$$
(19)

On the one hand, this allows to represent the random projection map onto \mathbb{R}^n using only O(nmd) parameters as opposed to the $O(nd^m)$ parameters in a densely parametrized random projection; and on the other, it allows for downstream computations to exploit the low-rank structure of the operator, as we shall do so in the definition stated below.

Definition 3.7. Let $\mathbf{W}^{(1)}, \dots, \mathbf{W}^{(M)} \stackrel{i.i.d.}{\sim} \Lambda^{\tilde{d}}$ be i.i.d. random matrices sampled from $\Lambda^{\tilde{d}}$ for RFF dimension $\tilde{d} \in \mathbb{Z}_+$, the define the independent RFF maps $\tilde{\varphi}_m : \mathfrak{X} \to \tilde{\mathcal{H}}$ as in (2), i.e. $\tilde{\varphi}_m(\mathbf{x}) = \frac{1}{\sqrt{d}} \left(\cos(\mathbf{W}^{(m)} \mathbf{x}), \sin(\mathbf{W}^{(m)} \mathbf{x}) \right)$ for $m \in [M]$ and $\mathbf{x} \in \mathfrak{X}$, and let $\mathbf{P}^{(1)}, \dots, \mathbf{P}^{(M)} \stackrel{i.i.d.}{\sim} \mathcal{N}^{\tilde{d}} \left(0, \mathbf{I}_{2\tilde{d}} \right)$ be random matrices with i.i.d. standard normal entries. The *M*-truncated Tensor Random Projected Random Fourier Signature Feature (RFSF-TRP) map $\tilde{\varphi}_{\mathrm{Sig}_{\leq M}}^{\mathrm{TRP}} \mathfrak{X}_{\mathrm{seq}} \to \tilde{\mathcal{H}}_{\mathrm{Sig}}^{\mathrm{TRP}} = \mathbb{R}^{M\tilde{d}}$ is defined for truncation level $M \in \mathbb{Z}_+$ and $\mathbf{x} \in \mathfrak{X}_{\mathrm{seq}}$ as

$$\tilde{\varphi}_{\mathrm{Sig}_{\leq M}}^{\mathrm{TRP}}(\mathbf{x}) \coloneqq \frac{1}{\sqrt{\tilde{d}}} \left(\left(\sum_{\mathbf{i} \in \Delta_m(\ell_{\mathbf{x}}-1)} \left\langle \mathbf{p}_q^{(1)}, \delta \tilde{\varphi}_1(\mathbf{x}_{i_1}) \right\rangle \cdots \left\langle \mathbf{p}_q^{(m)}, \delta \tilde{\varphi}_m(\mathbf{x}_{i_m}) \right\rangle \right)_{q=1}^{\tilde{d}} \right)_{m=0}^{M} = \frac{1}{\sqrt{\tilde{d}}} \left(\sum_{\mathbf{i} \in \Delta_m(\ell_{\mathbf{x}}-1)} \left(\mathbf{P}^{(1)^{\mathsf{T}}} \delta \tilde{\varphi}_1(\mathbf{x}_{i_1}) \right) \odot \cdots \odot \left(\mathbf{P}^{(m)^{\mathsf{T}}} \delta \tilde{\varphi}_m(\mathbf{x}_{i_m}) \right) \right)_{m=0}^{M}, \quad (20)$$

where $\mathbf{P}^{(m)} = \left(\mathbf{p}_{q}^{(m)}\right)_{q=1}^{\tilde{d}} \in \mathbb{R}^{2\tilde{d} \times \tilde{d}}$, and \odot denotes the Hadamard product⁶. The RFSF-TRP kernel $\tilde{k}_{\text{Sig}_{\leq M}}^{\text{TRP}} : \mathfrak{X}_{\text{seq}} \times \mathfrak{X}_{\text{seq}} \to \mathbb{R}$ can then be directly computed for sequences $\mathbf{x}, \mathbf{y} \in \mathfrak{X}_{\text{seq}}$ by

$$\begin{split} \tilde{\mathsf{k}}_{\mathrm{Sig}_{\leq M}}^{\mathrm{TRP}}(\mathbf{x},\mathbf{y}) &\coloneqq \left\langle \tilde{\varphi}_{\mathrm{Sig}_{\leq M}}^{\mathrm{TRP}}(\mathbf{x}), \tilde{\varphi}_{\mathrm{Sig}_{\leq M}}^{\mathrm{TRP}}(\mathbf{y}) \right\rangle_{\tilde{\mathcal{H}}_{\mathrm{Sig}}} = \sum_{m=0}^{M} \left\langle \tilde{\varphi}_{\mathrm{Sig}_{m}}^{\mathrm{TRP}}(\mathbf{x}), \tilde{\varphi}_{\mathrm{Sig}_{m}}^{\mathrm{TRP}}(\mathbf{y}) \right\rangle_{\tilde{\mathcal{H}}_{\otimes m}} \\ &= \sum_{m=0}^{M} \tilde{\mathsf{k}}_{\mathrm{Sig}_{m}}^{\mathrm{TRP}}(\mathbf{x},\mathbf{y}) = \frac{1}{\tilde{d}} \sum_{m=0}^{M} \sum_{q=1}^{\tilde{d}} \sum_{\substack{\mathbf{i} \in \Delta_{m}(\ell_{\mathbf{x}}-1) \\ \mathbf{j} \in \Delta_{m}(\ell_{\mathbf{y}}-1)}} \prod_{p=1}^{m} \left\langle \mathbf{p}_{q}^{(p)}, \delta \tilde{\varphi}_{p}(\mathbf{x}_{i_{p}}) \right\rangle \left\langle \mathbf{p}_{q}^{(p)}, \delta \tilde{\varphi}_{p}(\mathbf{y}_{j_{p}}) \right\rangle, (21) \end{split}$$

where we defined the level-m RFSF-TRP kernel $\tilde{k}_{\text{Sig}_m}^{\text{TRP}}$: $\mathfrak{X}_{\text{seq}} \times \mathfrak{X}_{\text{seq}} \to \mathbb{R}$ for $m \leq M$ as $\tilde{k}_{\text{Sig}_m}^{\text{TRP}}(\mathbf{x}, \mathbf{y}) \coloneqq \langle \tilde{\varphi}_{\text{Sig}_m}^{\text{TRP}}(\mathbf{x}), \tilde{\varphi}_{\text{Sig}_m}^{\text{TRP}}(\mathbf{y}) \rangle_{\tilde{\mathcal{H}}^{\otimes m}}$ with the convention that $\tilde{k}_{\text{Sig}_0}^{\text{TRP}} \equiv 1$.

We remark that (20) is equivalent to the TRP operator (19) applied to the RFSF map (13) by exploiting bilinearity of the inner product, and using that it factorizes over the tensor components, as described in (4). Then, the unbiasedness of (21) follows from the fact that the TRP operator is an isometry under expectation, which is applied to the RFSF tensor $\tilde{\varphi}_{\text{Sig}_m}$, therefore $\tilde{k}_{\text{Sig}_m}^{\text{TRP}}$ kernel is conditionally an unbiased estimator of \tilde{k}_{Sig_m} given the RFSF weights $\mathbf{W}^{(1)}, \ldots, \mathbf{W}^{(m)} \in \mathbb{R}^{d \times \tilde{d}}$. By the tower rule for expectations, $\tilde{k}_{\text{Sig}_m}^{\text{TRP}}$ is an unbiased estimator of k_{Sig_m} is an unbiased estimator of k_{Sig_m} . The approximation quality is then governed by two factors: (i) how well the TRP projected kernel $\tilde{k}_{\text{Sig}_m}^{\text{TRP}}$ approximates \tilde{k}_{Sig_m} ; (ii) the quality of the approximation of \tilde{k}_{Sig_m} with

⁶The Hadamard product stands for component-wise multiplication of the vectors $\mathbf{x}, \mathbf{y} \in \mathbb{R}^n$, $\mathbf{x} \odot \mathbf{y} = (x_i y_i)_{i=1}^n$

respect to k_{Sig_m} . Note that (ii) has already been discussed in Theorem 3.2. Here, we state the following theoretical result which quantifies (i). Combining these two results by means of triangle inequality and union bounding quantifies that $\tilde{k}_{Sig_m}^{TRP}$ is a good estimator of k_{Sig_m} .

Theorem 3.8. Let $k : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ be a continuous, bounded, translation-invariant kernel with spectral measure Λ , which satisfies (15). Then, the following bound holds for RFSF-TRP kernel for signature level $m \in \mathbb{Z}_+$ sequences $\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq}$ and $\epsilon > 0$

$$\mathbb{P}\left[\left|\tilde{\mathsf{K}}_{\text{Sig}_{m}}^{\text{TRP}}(\mathbf{x},\mathbf{y}) - \tilde{\mathsf{K}}_{\text{Sig}_{m}}(\mathbf{x},\mathbf{y})\right| \ge \epsilon\right] \le C_{d,\Lambda} \exp\left(-\left(\frac{m^{2}\tilde{d}^{\frac{1}{2m}}\epsilon^{\frac{1}{m}}}{2\sqrt{2}e^{3}R \|\mathbf{x}\|_{1-\text{var}}} \|\mathbf{y}\|_{1-\text{var}}}\right)^{\frac{1}{2}}\right), \quad (22)$$

where the absolute constant is defined as $C_{d,\Lambda} := 2\left(1 + \frac{S}{2R} + \frac{S^2}{4R^2}\right)^d$.

The proof is given in the supplement under Theorem C.12 utilizing the hypercontractivity of Gaussian polynomials [35] that is used to quantify the concentration of the TRP estimator. The concentration of the RFSF-TRP kernel is then governed by Theorems 3.2 and 3.8 combined. Together, they show that for smaller values of ϵ (i.e. the regime change as discussed below Theorem 3.2), the probability has a polynomial plus a sub-Gaussian tail, while for large ϵ , it has a $(\frac{1}{2m})$ -subexponential tail due to (22), and the dominant convergence rate with respect to \tilde{d} is $(\frac{1}{4m})$ -subexponential. This means that in terms of convergence, RFSF-TRP is the slowest among the 3 variations introduced so far. However, it is also the most efficient in terms of overall dimension, hence computational complexity as well, since $\tilde{\mathcal{H}}_{Sig}^{TRP} = \mathbb{R}^{M\tilde{d}}$. Remark 3.9 discusses the computational complexity in detail.

Remark 3.9. Algorithm 3 demonstrates the computation of the RFSF-TRP map $\tilde{\varphi}_{\text{Sig}_{\leq M}}^{\text{TRP}}$ given a dataset of sequences $\mathbf{X} = (\mathbf{x}_i)_{i=1}^N \subset \mathcal{X}_{\text{seq}}$. Counting the operations, here we can deduce that the algorithm has an $O(MNL\tilde{d}(d+\tilde{d}))$ computational complexity. This variation is also linear in L, the maximal sequence length, although it is quadratic in \tilde{d} .

4 Experiments

Time series classification. We perform multivariate time series classification experiments to investigate the performance of the scalable RFSF variants compared to the full-rank signature kernel and other quadratic time baseline kernels, and further, to demonstrate the scalability to large-scale datasets, where the quadratic sample complexity becomes prohibitive. We use support vector machine (SVM) [76] classification with for classifying multivariate time series on datasets of various sizes. For the quadratic time kernels, the dual SVM formulation is used, while for kernels with explicit feature representations, we use the primal formulation that has linear complexity in the size of the dataset $n \in \mathbb{Z}_+$ aiding in scalability to truly large-scale datasets. For each considered kernel/feature, we use a GPU-based implementation provided in the KSig library⁷. For large-scale experiments with the featurized kernels, linear SVM implementation is used from the cuML library [62], while the dual SVM on small-scale datasets uses the sklearn library [56]. For multi-class problems, we use the one-vs-one classification strategy. This study is also the largest scale comparison of various signature kernels to date which extends the datasets considered in [68]. The hardware used was 2 computer clusters equipped with overall 8 NVIDIA 3080 Ti GPUs.

⁷https://github.com/tgcsaba/KSig

Table 1: Computational complexities of kernels in our experiments; $N \in \mathbb{Z}_+$ is the number of time series, $L \in \mathbb{Z}_+$ is their length, $d \in \mathbb{Z}_+$ is their state-space dimension, $M \in \mathbb{Z}_+$ is the signature truncation level, $\tilde{d} \in \mathbb{Z}_+$ is the RF dimension, $W \in \mathbb{Z}_+$ is the warping length in RWS.

RFSF-DP	RFSF-TRP	KSig	KSigPDE	RWS	GAK	RBF	RFF
$O\left(NL\tilde{d}\left(Md+2^{M}\right)\right)$	$O\left(NLM\tilde{d}\left(d+\tilde{d}\right)\right)$	$O\left(N^2L^2(M+d)\right)$	$O(N^2L^2d)$	O(NLWd)	$O(N^2L^2d)$	$O(N^2Ld)$	$O(NLd\tilde{d})$

Methods. We compare the proposed variants RFSF-DP and RFSF-TRP to the baselines described here: (1) the *M*-truncated Signature Kernel [39] KSig formulated via the kernel trick, and is a quadratic time baseline; (2) the Signature-PDE Kernel [68] KSigPDE, which uses the 2nd-order PDE solver and also has quadratic complexity; (3) the Global Alignment Kernel [17] GAK, the previous SOTA time series kernel before the signature and can be related to it, see [39, Sec. 5]; (4) Random Warping Series [92] RWS, which produces features by DTW alignments between the input and randomly sampled time series; (5) the RBF kernel, which treats the whole time series as a vector of length \mathbb{R}^{Ld} , (6) Random Fourier Reatures [58] RFF, which also treats the time series as a long vector. The complexities are compared in Table 1.

Hyperparameter selection. For each dataset-kernel, we perform cross-validation to select the optimal hyperparameters that are optimized over the Cartesian product of the following options. For each method that requires a static kernel, we use the RBF kernel with bandwidth hyperparameter $\sigma > 0$. This is specified by a rescaled median heuristic, i.e.

$$\sigma = \alpha \operatorname{med}\left\{ \left\| \mathbf{x}_{i} - \mathbf{x}_{j}^{\prime} \right\|_{2} / 2 : i \in [\ell_{\mathbf{x}}], j \in [\ell_{\mathbf{x}^{\prime}}], \mathbf{x}, \mathbf{x}^{\prime} \in \mathbf{X} \right\}, \text{ for } \alpha > 0,$$

where α is chosen from $\alpha \in \{10^{-3}, \dots, 10^3\}$ on a logarithmic grid with 19 steps. For each kernel that is not normalized by default (i.e. the GAK and RBF kernels are normalized, the former is because without normalization it blows up), we select whether to normalize to unit norm in feature space via $k(\mathbf{x}, \mathbf{y}) \mapsto k(\mathbf{x}, \mathbf{y})/\sqrt{k(\mathbf{x}, \mathbf{x})k(\mathbf{y}, \mathbf{y})}$. The SVM hyperparameter C > 0is selected from $C \in \{10^0, 10^1, 10^2, 10^3, 10^4\}$. Further, motivated by previous work that investigates the effect of path augmentations in the context of signature methods [54], we chose 3 augmentations to cross-validate over. First is parametrization encoding, which gives the classifier the ability to remove the warping invariance of a given sequence kernel, adding the time index as an additional coordinate, i.e. for each time series in the dataset $\mathbf{x} \in \mathbf{X}$, we augment it via $\mathbf{x} = (\mathbf{x}_i)_{i=1}^{\ell_{\mathbf{x}}} \mapsto (\beta i/\ell_{\mathbf{x}}, \mathbf{x}_i)_{i=1}^{\ell_{\mathbf{x}}}$, where $\beta > 0$ is the parametrization intensity chosen from $\beta \in \{10^0, 10^1, 10^2, 10^3, 10^4\}$. The second augmentation is the basepoint encoding, the role of which is to remove the translation invariance of signature features. Note that when the static base kernel is chosen to be a nonlinear kernel other than the Euclidean inner product, the signature kernel is not completely translation-invariant due to the state-space nonlinearities, but it is close being that by the L-Lipschitz property in Lemma C.2 valid for of the static kernels considered in this work. The basepoint encoding adds an initial **0** step at the beginning of each time series, i.e. for $\mathbf{x} \in \mathbf{X}$, $\mathbf{x} =$ $(\mathbf{x}_1, \ldots, \mathbf{x}_{\ell_x}) \mapsto (\mathbf{0}, \mathbf{x}_1, \ldots, \mathbf{x}_{\ell_x})$. The third augmentation is the lead-lag map, which is defined as $\mathbf{x} = (\mathbf{x}_1, \dots, \mathbf{x}_{\ell_x}) \mapsto ((\mathbf{x}_1, \mathbf{x}_1), (\mathbf{x}_2, \mathbf{x}_1), (\mathbf{x}_2, \mathbf{x}_2), \dots, (\mathbf{x}_{\ell_x}, \mathbf{x}_{\ell_x-1}), (\mathbf{x}_{\ell_x}, \mathbf{x}_{\ell_x})).$ For the truncation-based signature kernels, we select the truncation level $M \in \mathbb{Z}_+$ from $M \in \{2, 3, 4, 5\}$. For RWS, we select the warping length from $W \in \{10, 20, \dots, 100\}$ as suggested by the authors. This makes RWS the most expensive feature-based kernel, and so as to fit within the same resource limitations, we omit cross-validating over the path augmentations. For RF approaches, we set the RF dimension $\tilde{d} \in \mathbb{Z}_+$, so the overall dimension is 1000.

Datasets: UEA Archive. The UEA archive [18] is a collection of overall 30 datasets for benchmarking classifiers on multivariate time series classification problems containing both binary and multi-class tasks. The data modality ranges from various sources e.g. human activity recognition, motion classification, ECG classification, EEG/MEG classification, audio spectra recognition, and others. The sizes of the datasets in terms of number of time series range from small (\leq 1000 examples) to large (\leq 30000), and includes various lengths between 8 and 18000. A summary of the dataset characteristics can be found in Table 2 in [18]. Pre-specified train-test splits are provided for each dataset, which we follow. We evaluate all considered kernels on the small datasets (\leq 1000 time series), but because the non-feature-based become very expensive computationally beyond these sizes, we only evaluate feature-based approaches on medium and large datasets (\geq 1000 time series). Each featurized approach is trained and evaluated 5 times on each dataset in order to account for the randomness in the hyperparameter selection procedure and evaluation.

Datasets: Mental Workload Intensity Classification. We evaluate featurized approaches on a large-scale brain-activity recording data set called fNIRS2MW.⁸ This dataset contains brain activity recordings collected from overall 68 participants during a 30-60 minute experimental session, where they were asked to carry out tasks of varying intensity. The collected time series are sliced into 30 second segments using a sliding window, and each segment is labelled with an intensity level (0-3), giving rise to overall ~ 100000 segments, which we split in a ratio of 80-20 for training and testing. We convert the task into a binary classification problem by assigning a label whether the task is low (0 or 1) or high (2 or 3) intensity.

Datasets: Satellite Image Classification. As a massive scale task, we use a satellite imagery dataset⁹ of $N = 10^6$ time series. Each length L = 46 time series corresponds to a vegetation index calculated from remote sensing data, and the task is to classify land cover types [57] by mapping vegetation profiles to various types of crops and forested areas corresponding to 24 classes. We split the dataset in a ratio of 90-10 for training and testing.

4.1 Results

Table 2 compares test accuracies on moderate size multivariate time series classification datasets with $N \leq 1000$ from the UEA archive. KSig provides state-of-the-art performance among all sequence kernels with taking the highest aggregate score in terms of all of average accuracy, average rank, and number of first places. Our proposed random feature variants RFSF-DP and RFSF-TRP provide comparable performance on most of the datasets in terms of accuracy, and they are only outperformed by KSig and KSigPDE with respect to average accuracy and rank. Interestingly, RFSF-TRP has more first place rankings, but RFSF-DP performs slightly better on average. This shows that on datasets of these sizes, using either of RFSF-DP and RFSF-TRP does not sacrifice model performance - even leading to improvements in some cases, potentially due to the implicit regularization effect of restricting to a finite-dimensional feature space - and it can already provide speedups. We visualize the critical difference diagram comparison of all considered approaches in Figure 1.

Table 3 demonstrates the performance of scalable approaches, i.e. RFSF-DP, RFSF-TRP, RWS and RFF on the remaining UEA datasets ($N \ge 1000$), the dataset fNIRS2MW ($N = 10^5$), and the satellite dataset SITS1M ($N = 10^6$). We find it infeasible to perform full cross-validation

⁸https://github.com/tufts-ml/fNIRS-mental-workload-classifiers

⁹https://cloudstor.aarnet.edu.au/plus/index.php/s/pRLVtQyNhxDdCoM

	RFSF-DP	RFSF-TRP	KSig	KSigPDE	RWS	GAK	RBF	RFF
ArticularyWordRecognition	0.984	0.981	0.990	0.983	0.987	0.977	0.977	0.978
AtrialFibrillation	0.373	0.320	0.400	0.333	0.427	0.333	0.267	0.373
BasicMotions	1.000	1.000	1.000	1.000	0.995	1.000	0.975	0.860
Cricket	0.964	0.964	0.958	0.972	0.978	0.944	0.917	0.886
DuckDuckGeese	0.636	0.664	0.700	0.480	0.492	0.500	0.420	0.372
ERing	0.921	0.936	0.841	0.941	0.945	0.926	0.937	0.915
EigenWorms	0.817	0.837	0.809	0.794	0.623	0.511	0.496	0.443
Epilepsy	0.949	0.942	0.949	0.891	0.925	0.870	0.891	0.777
EthanolConcentration	0.457	0.439	0.479	0.460	0.284	0.361	0.346	0.325
FingerMovements	0.608	0.624	0.640	0.630	0.612	0.500	0.620	0.570
HandMovementDirection	0.573	0.568	0.595	0.527	0.403	0.595	0.541	0.454
Handwriting	0.434	0.400	0.479	0.409	0.591	0.481	0.307	0.249
Heartbeat	0.717	0.712	0.712	0.722	0.714	0.717	0.717	0.721
JapaneseVowels	0.978	0.978	0.986	0.986	0.955	0.981	0.981	0.979
Libras	0.898	0.928	0.922	0.894	0.837	0.767	0.800	0.800
MotorImagery	0.516	0.526	0.500	0.500	0.508	0.470	0.500	0.482
NATOPS	0.906	0.908	0.922	0.928	0.924	0.922	0.917	0.900
PEMS-SF	0.800	0.808	0.827	0.838	0.701	0.855	0.855	0.770
RacketSports	0.874	0.861	0.921	0.908	0.878	0.849	0.809	0.755
SelfRegulationSCP1	0.868	0.856	0.904	0.904	0.829	0.915	0.898	0.885
SelfRegulationSCP2	0.489	0.510	0.539	0.544	0.481	0.511	0.439	0.492
StandWalkJump	0.387	0.333	0.400	0.400	0.347	0.267	0.533	0.267
UWaveGestureLibrary	0.882	0.881	0.912	0.866	0.897	0.887	0.766	0.846
Avg. acc.	0.740	0.738	0.756	0.735	0.710	0.702	0.692	0.656
Avg. rank	3.609	3.739	2.348	2.957	3.957	4.174	4.913	5.913

Table 2: Comparison of SVM test accuracies on small multivariate time series classification datasets. For each row, the best result is highlighted in **bold**, and the second best in *italic*.



Figure 1: Critical difference diagram comparison on moderate datasets of considered approaches using two-tailed Nemenyi test [19].

Table 3: Comparison of accuracies on largescale datasets of random features.

	RFSF-DP	RFSF-TRP	RWS	RFF
CharacterTrajectories	0.990	0.990	0.991	0.989
FaceDetection	0.653	0.656	0.642	0.572
InsectWingbeat	0.436	0.459	0.227	0.341
LSST	0.589	0.624	0.631	0.423
PenDigits	0.983	0.982	0.989	0.980
PhonemeSpectra	0.204	0.204	0.205	0.083
SITS1M	0.745	0.740	0.610	0.718
SpokenArabicDigits	0.981	0.980	0.981	0.964
fNIRS2MW	0.659	0.658	0.621	0.642
Avg. acc.	0.693	0.699	0.655	0.635
Avg. rank	1.778	1.889	2.222	3.333

for quadratic time kernels on these datasets due to expensive kernel computations and downstream cost of dual SVM. The results show that both variants RFSF-DP and RFSF-TRP perform significantly better on average with respect to accuracy and rank then both RWS and RFF. Note when RWS takes first place, it only improves over our approach marginally, however, when it underperforms, it often does so severely. This is not surprising as both RFSF-DP and RFSF-TRP approximate the signature kernel, which is universal kernel on time series; it is theoretically capable of learning from any kind of time series data as supported by its best overall performance above.

5 Conclusion

We constructed a random kernel $\tilde{k}_{Sig_{\leq M}}$ for sequences that benefits from (i) lifting the original sequence to an infinite-dimensional RKHS \mathcal{H} , (ii) linear complexity in sequence length, (iii) being with high probability close the signature kernel k_{Sig} . Thereby it combines the strength of the signature kernel k_{Sig} which is to implicitly use the iterated integrals of a sequence that has an infinite-dimensional RKHS \mathcal{H} as state-space with the strength of (unkernelized) signature features φ_{Sig} that only require linear time complexity. Our main theoretical result extends the theoretical guarantees for translation-invariant kernels on linear spaces to the signature kernel defined on the nonlinear domain \mathcal{X}_{seq} ; however, the proofs differ from the classic case and require to analyse the error propagation in tensor space. A second step is more straightforward, and combines this approach with random projections in finite-dimensions for tensors to reduce the complexity in memory further. The advantages and disadvantages of the resulting approach are analogous to the classic RFF technique on \mathbb{R}^d , namely a reduction of computational complexity by an order for the price of an approximation that only holds with high probability. As in the classic RFF case, our experiments indicate that this is in general a favourable tradeoff.

In the future, it would be interesting both theoretically and empirically to replace the vanilla Monte Carlo integration in the RFF construction by block-orthogonal random matrices as done in [94]. Further, our random features can also be used to define an unbiased appoximation to the inner product of expected signatures, which has found usecases, among many, in nonparametric hypothesis testing and market regime detection [10, 32], training of generative models [55, 33], and graph representation learning [83].

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A Concentration of measure

Classic inequalities. The following inequalities are classic, and their proofs are in analysis textbooks, e.g. [21, 66]. Firstly, Jensen's inequality is useful for convex (concave) expectations.

Lemma A.1 (Jensen's inequality). Let X be an integrable random variable, and $f : \mathbb{R} \to \mathbb{R}$ a convex function, such that f(X) is also integrable. Then, the following inequality holds:

$$f(\mathbb{E}[X]) \le \mathbb{E}[f(X)].$$

Hölder's inequality is a generalization of the Cauchy-Schwarz inequality to L^p spaces.

Lemma A.2 (Hölder's inequality). Let $p, q \ge 1$ such that $\frac{1}{p} + \frac{1}{q} = 1$. Let X and Y respectively be L^p and L^q integrable random variables, i.e. $\mathbb{E}[|X|^p] < \infty$ and $\mathbb{E}[|Y|^q] < \infty$. Then, XY is integrable, and it holds that

$$\mathbb{E}[|XY|] \le \mathbb{E}^{1/p} \left[|X|^p \right] \mathbb{E}^{1/q} \left[|Y|^q \right].$$

Although not inherently a probabilistic inequality, Young's inequality can be used to decouple products of random variables.

Lemma A.3 (Young's inequality). Let p, q > 0 with $\frac{1}{p} + \frac{1}{q} = 1$. Then, for every $a, b \ge 0$

$$ab \le \frac{a^p}{p} + \frac{b^q}{q}$$

Lemma A.4 (Reverse Young's inequality). Let p, q > 0 such that $\frac{1}{p} - \frac{1}{q} = 1$. Then, for every $a \ge 0$ and b > 0, it holds that

$$ab \ge \frac{a^p}{p} - \frac{b^{-q}}{q}$$

Proof. Apply Young's inequality with $p' = \frac{1}{p}$ and $q' = \frac{q}{p}$ to $a' = (ab)^p$ and $b' = b^{-p}$.

Subexponential concentration. Next, we state a variation on the well-known Bernstein inequality, which holds for random variables in the subexponential class. The condition (25) on a random variable *X*, in this case, is formulated as a moment-growth bound, called a Bernstein moment condition. We show in Lemmas A.11 and A.12 (in a more general setting) that (25) is equivalent to the random variable being subexponential. Further, note that the condition itself is given in terms of non-centered random variables, while the statement of the theorem is about their centered counterparts. For similar results, see [86, Sec. 2.2.2].

Theorem A.5 (Bernstein inequality - one-tailed). Let $X_1, ..., X_n$ be independent random variables satisfying the following moment-growth condition for some S, R > 0,

$$\mathbb{E}\left[X_{i}^{k}\right] \leq \frac{k!S^{2}R^{k-2}}{2} \quad for \ all \ k \geq 2.$$

$$(23)$$

Then, it holds for $\tilde{X}_i := X_i - \mathbb{E}[X_i]$ *that*

$$\mathbb{P}\left[\sum_{i=1}^{n} \tilde{X}_{i} \ge t\right] \le \exp\left(\frac{-t^{2}}{2(nS^{2}+Rt)}\right)$$

Proof. We have for $\lambda > 0$

$$\mathbb{P}\left[\sum_{i=1}^{n} \tilde{X}_{i} \geq t\right]^{(a)} \leq \exp(-\lambda t) \mathbb{E}\left[\exp\left(\lambda \sum_{i=1}^{n} \tilde{X}_{i}\right)\right]$$
$$\stackrel{(b)}{=} \exp\left(-\lambda t - \lambda \sum_{i=1}^{n} \mathbb{E}[X_{i}]\right) \prod_{i=1}^{n} \mathbb{E}[\exp(\lambda X_{i})], \quad (24)$$

where (a) holds for any $\lambda > 0$ by the Chernoff bound applied to $\sum_{i=1}^{n} \tilde{X}_i$, and in (b) we used the independence of X_i -s. Bounding the moment-generating function of X_i , one gets

$$\mathbb{E}[\exp(\lambda X_i)] \stackrel{\text{(d)}}{=} 1 + \lambda \mathbb{E}[X_i] + \sum_{k=2}^{\infty} \frac{\lambda^k}{k!} \mathbb{E}[X_i^k] \stackrel{\text{(e)}}{\leq} 1 + \lambda \mathbb{E}[X_i] + \frac{S^2 \lambda^2}{2} \sum_{k=0}^{\infty} \lambda^k R^k$$
$$\stackrel{\text{(f)}}{=} 1 + \lambda \mathbb{E}[X_i] + \frac{S^2 \lambda^2}{2(1 - \lambda R)} \stackrel{\text{(g)}}{\leq} \exp\left(\lambda \mathbb{E}[X_i] + \frac{S^2 \lambda^2}{2(1 - \lambda R)}\right),$$

where (d) is due to the Taylor expansion of the exponential function, (e) implied by (23), (f) holds by the sum of geometric series for any $\lambda < 1/R$, (g) follows from the inequality $1 + x \le \exp(x)$ for all $x \in \mathbb{R}$. Using this to bound (24),

$$\begin{split} \mathbb{P}\Big[\sum_{i=1}^{n} \tilde{X}_{i} \geq t\Big] &\leq \exp\Big(-\lambda t - \lambda \sum_{i=1}^{n} \mathbb{E}[X_{i}]\Big) \prod_{i=1}^{n} \exp\Bigg(\lambda \mathbb{E}[X_{i}] + \frac{S^{2}\lambda^{2}}{2(1-\lambda R)}\Bigg) \\ &= \exp\Bigg(-\lambda t + \frac{nS^{2}\lambda^{2}}{2(1-\lambda R)}\Bigg). \end{split}$$

Finally, choosing $\lambda = t/(nS^2 + Rt)$ gives the statement.

Corollary A.6 (Bernstein inequality - two-tailed). Let $X_1, ..., X_n$ be independent random variables satisfying the the following moment-growth condition for some S, R > 0,

$$\mathbb{E}\left[\left|X_{i}\right|^{k}\right] \leq \frac{k!S^{2}R^{k-2}}{2} \quad for \ all \ k \geq 2.$$

$$(25)$$

Then, it holds for $\tilde{X}_i := X_i - \mathbb{E}[X_i]$ *that*

$$\mathbb{P}\left[\left|\sum_{i=1}^{n} \tilde{X}_{i}\right| \geq t\right] \leq 2 \exp\left(\frac{-t^{2}}{2(nS^{2}+Rt)}\right).$$

Proof. Applying Theorem A.5 to $\sum_{i=1}^{n} \tilde{X}_i$ and $-\sum_{i=1}^{n} \tilde{X}_i$, and combining the two bounds gives the two-tailed result.

 α -exponential concentration. In this section, we introduce a specific class of Orlicz norms for heavy-tailed random variables, which generalizes subexponential distributions.

Definition A.7 (α -exponential Orlicz norm). Let $\alpha > 0$ and define the function

$$\Psi_{\alpha}: \mathbb{R} \to \mathbb{R}, \quad \Psi_{\alpha}(x) \coloneqq \exp\left(x^{\alpha}\right) - 1 \quad for \ all \ x \in \mathbb{R}.$$

The α -exponential Orlicz (quasi-)norm of a random variable X is defined as

$$\|X\|_{\Psi_{\alpha}} \coloneqq \inf\left\{t > 0 : \mathbb{E}\left[\Psi_{\alpha}\left(\frac{|X|}{t}\right)\right] \le 1\right\},\$$

adhering to the standard convention that $\inf \emptyset = \infty$.

If a random variable *X* satisfies $||X||_{\Psi_{\alpha}} < \infty$, it is either called an α -(sub)exponential random variable (if $\alpha < 1$) [69, 30, 9], or sub-Weibull of order- α [41, 95].

An alternative characterization of the α -exponential norm is the following tail-bound.

Remark A.8 (Tail bound). Let $\alpha > 0$ and X be a random variable such that $||X||_{\Psi_{\alpha}} < \infty$. *Then,*

$$\mathbb{P}[|X| \ge \epsilon] \le 2 \exp\left(-\frac{\epsilon^{\alpha}}{\|X\|_{\Psi_{\alpha}}^{\alpha}}\right).$$

Proof. Due to Definition A.7, $\mathbb{E}\left[\exp\left(\frac{|X|^{\alpha}}{\|X\|_{\Psi_{\alpha}}^{\alpha}}\right)\right] \le 2$, hence by Markov's inequality

$$\mathbb{P}[|X| \ge \epsilon] = \mathbb{P}\left[\frac{|X|^{\alpha}}{\|X\|_{\Psi_{\alpha}}^{\alpha}} \ge \frac{\epsilon^{\alpha}}{\|X\|_{\Psi_{\alpha}}^{\alpha}}\right] \le e^{-\frac{\epsilon^{\alpha}}{\|X\|_{\Psi_{\alpha}}^{\alpha}}} \mathbb{E}\left[\exp\left(\frac{|X|^{\alpha}}{\|X\|_{\Psi_{\alpha}}^{\alpha}}\right)\right] \le 2e^{-\frac{\epsilon^{\alpha}}{\|X\|_{\Psi_{\alpha}}^{\alpha}}}.$$

Note that although $\|\cdot\|_{\Psi_{\alpha}}$ is often referred to as a norm, it does not satisfy the triangle inequality for $\alpha < 1$, although we can still relate the norm of the sum to the sum of the norms.

Lemma A.9 (Generalized triangle inequality for Orlicz norm). *It holds for any random variables X*, *Y* and $\alpha > 0$ that

$$||X + Y||_{\Psi_{\alpha}} \le C_{\alpha} (||X||_{\Psi_{\alpha}} + ||Y||_{\Psi_{\alpha}}),$$

where $C_{\alpha} = 2^{1/\alpha}$ if $\alpha < 1$ and 1 otherwise.

Proof. See Lemma A.3. in [30].

A useful property of α -exponential norms is that they satisfy a Hölder-type inequality.

Lemma A.10 (Hölder inequality for Orlicz norm). *It holds for any random variables* X_1, \ldots, X_k and $\alpha_1, \ldots, \alpha_k > 0$ that

$$\left\|\prod_{i=1}^{k} X_{i}\right\|_{\Psi_{a}} \leq \prod_{i=1}^{k} \|X_{i}\|_{\Psi_{a_{i}}},$$

where $\alpha \coloneqq \left(\sum_{i=1}^k \alpha_i^{-1}\right)^{-1}$.

Proof. See Lemma A.1. in [30].

Next, we show in the following two lemmas that a random variable *X* is α -exponential if and only if $|X|^{\alpha}$ satisfies a Bernstein moment-growth condition.

Lemma A.11 (Bernstein condition implies Orlicz norm bound). Let $\alpha > 0$ and X be a random variable that satisfies for S, R > 0 that

$$\mathbb{E}\left[|X|^{k\alpha}\right] \le \frac{k! S^2 R^{k-2}}{2} \quad \text{for all } k \ge 2.$$
(26)

Then, it holds that

$$||X||_{\Psi_{\alpha}} \leq (S \vee R)^{1/\alpha}$$

Proof. Firstly, due to Jensen's inequality (Lemma A.1) and (26), we have

$$\mathbb{E}\left[|X|^{\alpha}\right] \le \mathbb{E}^{1/2}\left[|X|^{2\alpha}\right] \le S.$$
(27)

We proceed similarly to the proof of Theorem A.5. For t > 0, we have

$$\mathbb{E}\left[\exp\frac{|X|^{\alpha}}{t^{\alpha}}\right] \stackrel{(a)}{=} 1 + \frac{\mathbb{E}\left[|X|^{\alpha}\right]}{t^{\alpha}} + \sum_{m=2}^{\infty} \frac{\mathbb{E}\left[|X|^{m\alpha}\right]}{t^{m\alpha}m!} \stackrel{(b)}{\leq} 1 + \frac{S}{t^{\alpha}} + \frac{S^{2}}{2t^{2\alpha}} \sum_{m=0}^{\infty} \left(\frac{R}{t^{\alpha}}\right)^{m}$$
$$\stackrel{(c)}{=} \underbrace{1 + \frac{S}{t^{\alpha}} + \frac{S^{2}}{2t^{\frac{1}{2}\alpha}} \frac{t^{\frac{1}{2}\alpha}}{t^{\alpha} - R}}_{f(t)},$$

where (a) is due to the Taylor expansion of the exponential function, (b) is due to (26) and (27), (c) is the sum of a geometric series for $R < t^{\alpha}$. Defining $f(t) \coloneqq 1 + \frac{S}{t^{\alpha}} \left(1 + \frac{S}{2(t^{\alpha} - R)}\right)$ and solving for $f(t) \le 2$ leads to the quadratic inequality

$$0 \le t^{2\alpha} - (S+R)t^{\alpha} + SR - \frac{S^2}{2},$$

which has roots $t_{1,2}^{\alpha} = \frac{1}{2} \left(S + R \pm \sqrt{(S - R)^2 + 2S^2} \right)$. We discard the left branch, which violates the condition $R < t^{\alpha}$ since

$$t_{2}^{\alpha} = \frac{1}{2} \left(S + R - \sqrt{(S - R)^{2} + 2S^{2}} \right) \le \frac{1}{2} \left(S + R - |S - R| \right) = S \land R \le R.$$

Now, as the inequality $\sqrt{x^2 + y^2} \le |x| + |y|$ holds for all $x, y \in \mathbb{R}$, we get

$$\begin{split} t_1^{\alpha} &\leq \frac{1}{2} \Big(S + R + \sqrt{(S - R)^2 + 2S^2} \Big) \leq \frac{1}{2} \Big(S + R + |S - R| + \sqrt{2}S \Big) \\ &= S \lor R + \frac{\sqrt{2}}{2} S \leq 2(S \lor R) \end{split}$$

and hence choosing $t \ge (2(S \lor R))^{1/\alpha} \ge t_1$ implies $f(t) \le 2$, and we are done.

The other direction is proven in the following lemma.

Lemma A.12 (Finite Orlicz norm implies Bernstein condition). Let $\alpha > 0$ and X be a random variable such that $||X||_{\Psi_{\alpha}} < \infty$. Then,

$$\mathbb{E}\left[|X|^{k\alpha}\right] \le \frac{k! S^2 R^{k-2}}{2} \quad \text{for all } k \ge 2,$$

where $S \coloneqq \sqrt{2} \|X\|_{\Psi_q}^{\alpha}$ and $R \coloneqq \|X\|_{\Psi_q}^{\alpha}$.

Proof. We have that

$$\mathbb{E}\left[|X|^{k\alpha}\right] \stackrel{\text{(a)}}{=} k! \|X\|_{\Psi_{\alpha}}^{k\alpha} \mathbb{E}\left[\frac{1}{k!} \left(\frac{|X|}{\|X\|_{\Psi_{\alpha}}}\right)^{k\alpha}\right] \stackrel{\text{(b)}}{\leq} k! \|X\|_{\Psi_{\alpha}}^{k\alpha} \mathbb{E}\left[\exp\left(\frac{|X|^{\alpha}}{\|X\|_{\Psi_{\alpha}}^{\alpha}}\right) - 1\right]$$
$$\stackrel{\text{(c)}}{\leq} k! \|X\|_{\Psi_{\alpha}}^{k\alpha} \stackrel{\text{(d)}}{=} \frac{k! \left(\sqrt{2} \|X\|_{\Psi_{\alpha}}^{\alpha}\right)^{2} \left(\|X\|_{\Psi_{\alpha}}^{\alpha}\right)^{k-2}}{2},$$

where (a) is simply multiplying and dividing by the same values, (b) is the inequality $1 + x^k/k! \le e^x$ for x > 0, (c) follows from Definition A.7, (d) is reorganizing terms to the required form.

We adapt the following concentration inequality from [41] for α -exponential summation.

Theorem A.13 (Concentration inequality for α -subexponential summation). Let $\alpha \in (0, 1)$ and X_1, \ldots, X_n be independent, centered random variables with $||X_i||_{\Psi_{\alpha}} \leq M_{\alpha}$ for all $i \in [n]$. Then, it holds for t > 0 that

$$\mathbb{P}\left[\left|\sum_{i=1}^{n} X_{i}\right| \geq C_{\alpha}\left(2\sqrt{nt} + \sqrt{2}4^{1/\alpha}t^{1/\alpha}\right)\right] \leq 2e^{-t},\tag{28}$$

where $C_{\alpha} \coloneqq \sqrt{8}e^{4}(2\pi)^{1/4}e^{1/24}(2e/\alpha)^{1/\alpha}M_{\alpha}$. Alternatively, it holds for $\epsilon > 0$ that

$$\mathbb{P}\Big[\Big|\sum_{i=1}^{n} X_i\Big| \ge \epsilon\Big] \le 2\exp\left(-\frac{1}{4}\min\left\{\left(\frac{\epsilon}{2\sqrt{n}C_{\alpha}}\right)^2, \left(\frac{\epsilon}{\sqrt{8}C_{\alpha}}\right)^{\alpha}\right\}\right).$$
(29)

Proof. The inequality (28) follows directly from [41, Theorem 3.1].

To show (29), let $g_1(t) \coloneqq 2C_\alpha \sqrt{nt}$ and $g_2(t) \coloneqq \sqrt{2}C_\alpha 4^{1/\alpha} t^{1/\alpha}$. Now, for t > 0

$$\mathbb{P}\Big[\Big|\sum_{i=1}^{n} X_{i}\Big| \ge 2\big(g_{1}(t) \lor g_{2}(t)\big)\Big] \le \mathbb{P}\Big[\Big|\sum_{i=1}^{n} X_{i}\Big| \ge g_{1}(t) + g_{2}(t)\Big] \le 2e^{-t},$$

which is equivalently written as

$$\mathbb{P}\left[\left|\sum_{i=1}^{n} X\right| \ge \epsilon\right] \le 2 \exp\left(-\left(g_1^{-1}(\epsilon/2) \wedge g_2^{-1}(\epsilon/2)\right)\right).$$

Hypercontractivity. Here we provide an alternative approach for the concentration of heavy-tailed random variables, specifically for polynomials of Gaussian random variables. The following lemma states a moment bound for such Gaussian chaoses, considered in e.g. [35, 4].

Lemma A.14 (Moment bound for Gaussian polynomial). *Consider a degree-p polynomial* $f(X) = f(X_1, ..., X_n)$ of independent centered Gaussian random variables, $X_1, ..., X_n \stackrel{i.i.d.}{\sim} \mathcal{N}(0, 1)$. *Then, for all* $k \ge 2$

$$\mathbb{E}^{1/k} \Big[\big| f(X) \big|^k \Big] \le (k-1)^{p/2} \mathbb{E}^{1/2} \Big[\big| f(X) \big|^2 \Big].$$

Theorem A.15 (Concentration inequality for Gaussian polynomial). Consider a degree-*p* polynomial $f(X) = f(X_1, ..., X_n)$ of independent centered Gaussian random variables, $X_1, ..., X_n \stackrel{i.i.d.}{\sim} \mathcal{N}(0, 1)$. Then, for $p \ge 2$ and $\epsilon > 0$

$$\mathbb{P}\left[\left|f(X) - \mathbb{E}\left[f(X)\right]\right| \ge \epsilon\right] \le 2 \exp\left(-\frac{\epsilon^{2/p}}{2\sqrt{2}e^{\mathbb{V}^{1/p}}\left[f(X)\right]}\right),$$

where $\mathbb{V}[\cdot]$ *denotes the variance of a random variable.*

Proof. Without loss of generality, we may assume that $\mathbb{E}[f(X)] = 0$ and $\mathbb{V}[f(X)] = 1$. Since Lemma A.14 holds, we have for $p, k \ge 2$

$$\mathbb{E}\left[\left|f(X)\right|^{2k/p}\right] \stackrel{\text{(a)}}{\leq} \mathbb{E}^{2/p}\left[\left|f(X)\right|^{k}\right] \stackrel{\text{(b)}}{\leq} (k-1)^{k} \le k^{k} \stackrel{\text{(c)}}{\leq} k! e^{k} \le \frac{k! (\sqrt{2}e)^{2} e^{k-2}}{2}$$

where (a) holds due to Jensen inequality since $(\cdot)^{2/p}$ is concave, (b) is Lemma A.14, and (c) is due to Stirling's approximation. Hence, $|f(X)|^{2/p}$ satisfies a Bernstein condition with $S = \sqrt{2}e$ and R = e. Therefore, by Lemma A.11, we have $||f(X)||_{\Psi_{2/p}} \leq (2\sqrt{2}e)^{p/2}$.

Then, by Remark A.8, it holds that

$$\mathbb{P}\left[\left|f(X)\right| \ge \epsilon\right] \le 2 \exp\left(-\frac{\epsilon^{2/p}}{\left\|f(X)\right\|_{\Psi_{2/p}}^{2/p}}\right) \le 2 \exp\left(-\frac{\epsilon^{2/p}}{2\sqrt{2}e}\right).$$

B Random Fourier Features

[58] provides a uniform bound for the RFF error over a compact and convex domain $\mathcal{X} \subset \mathbb{R}^d$ with diameter $|\mathcal{X}|$ for some absolute constant C > 0,

$$\mathbb{P}\left[\sup_{\mathbf{x},\mathbf{y}\in\mathcal{X}}\left|\tilde{\mathsf{k}}(\mathbf{x},\mathbf{y})-\mathsf{k}(\mathbf{x},\mathbf{y})\right| \ge \epsilon\right] \le C\left(\frac{\sigma_{\Lambda}|\mathcal{X}|}{\epsilon}\right)^2 \exp\left(-\frac{-\tilde{d}\epsilon^2}{4(d+2)}\right),\tag{30}$$

where $\sigma_{\Lambda}^2 = \mathbb{E}_{\mathbf{w} \sim \Lambda} [\|\mathbf{w}\|^2]$, and [79] shows that $C \leq 66$. Equation (30) implies [74, Sec. 2]

$$\sup_{\mathbf{x},\mathbf{y}\in\mathcal{X}} \left| \tilde{\mathsf{k}}(\mathbf{x},\mathbf{y}) - \mathsf{k}(\mathbf{x},\mathbf{y}) \right| = O_p \left(|\mathcal{X}| \sqrt{\tilde{d}^{-1}\log \tilde{d}} \right).$$

The converse result¹⁰ for RFF derivatives was shown in [74, Thm. 5]. The idea of the proof is to cover the input domain by an ϵ -net, and control the approximation error on the centers, while simultaneously controlling the Lipschitz constant of the error to get the bound to hold uniformly. We provide an adapted version with two main differences: (1) using the Bernstein inequality from Corollary A.6, where the Bernstein condition is given in terms of non-centered random variables, and (2) using the covering numbers of [16]. We will use this theorem in proving Theorem 3.2 for controlling the approximation error of the derivatives of RFFs.

Given $\mathbf{p} \in \mathbb{N}^d$, we denote $|\mathbf{p}| \coloneqq p_1 + \ldots + p_d$, for a function $f : \mathbb{R}^d \to \mathbb{R}$ the **p**-th partial derivative by $\partial^{\mathbf{p}} f(\mathbf{z}) \coloneqq \frac{\partial^{|\mathbf{p}|} f(\mathbf{z})}{\partial^{p_1} z_1 \ldots \partial^{p_d} z_d}$, for a vector $\mathbf{w} \in \mathbb{R}^d$ the **p**-th power by $\mathbf{w}^{\mathbf{p}} \coloneqq w_1^{p_1} \cdots w_d^{p_d}$.

Theorem B.1 (Concentration inequality for RFF kernel derivatives). Let $\mathbf{p}, \mathbf{q} \in \mathbb{N}^d$ and $\mathbf{k} : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ be a continuous, bounded, translation-invariant kernel such that $\mathbf{z} \to \nabla[\partial^{\mathbf{p},\mathbf{q}}\mathbf{k}(\mathbf{z})]$ is continuous. Let $\mathcal{X} \subset \mathbb{R}^d$ be a compact and convex domain with diameter $|\mathcal{X}|$, and denote by $D_{\mathbf{p},\mathbf{q},\mathcal{X}} \coloneqq \sup_{\mathbf{z}\in\mathcal{X}_\Delta} ||\nabla[\partial^{\mathbf{p},\mathbf{q}}\mathbf{k}(\mathbf{z})]||_2$, where $\mathcal{X}_\Delta \coloneqq \{\mathbf{x}-\mathbf{y}: \mathbf{x}, \mathbf{y}\in\mathcal{X}\}$. Let Λ be the spectral measure of \mathbf{k} satisfying that $E_{\mathbf{p},\mathbf{q}} \coloneqq \mathbb{E}_{\mathbf{w}\sim\Lambda}[|\mathbf{w}^{\mathbf{p}+\mathbf{q}}| ||\mathbf{w}||_2] < \infty$, and the Bernstein moment condition for some S, R > 0,

$$\mathbb{E}_{\mathbf{w} \sim \Lambda} \Big[\left| \mathbf{w}^{\mathbf{p} + \mathbf{q}} \right|^k \Big] \leq \frac{k! S^2 R^{k-2}}{2} \quad for \ all \ k \geq 2.$$

¹⁰These error guarantees can also be improved exponentially both for the approximation of kernel values [74] and kernel derivatives [80, 9] in terms of the size of the domain where it holds, i.e. |X|.

Then, for $C_{\mathbf{p},\mathbf{q},\mathfrak{X}} := |\mathfrak{X}| (D_{\mathbf{p},\mathbf{q},\mathfrak{X}} + E_{\mathbf{p},\mathbf{q}})$ and $\epsilon > 0$,

$$\mathbb{P}\left[\sup_{\mathbf{x},\mathbf{y}\in\mathcal{X}} \left|\partial^{\mathbf{p},\mathbf{q}}\tilde{\mathbf{k}}(\mathbf{x},\mathbf{y}) - \partial^{\mathbf{p},\mathbf{q}}\mathbf{k}(\mathbf{x},\mathbf{y})\right| \ge \epsilon\right] \le 16\left(\frac{C_{\mathbf{p},\mathbf{q},\mathcal{X}}}{\epsilon}\right)^{\frac{d}{d+1}} \exp\left(\frac{-\tilde{d}\epsilon^2}{4(d+1)(2S^2 + R\epsilon)}\right).$$

Proof. We adapt the proof of [58, 74]. Note that as \mathcal{X} is compact, so is \mathcal{X}_{Δ} , and it can be covered by an ϵ -net of at most $T \coloneqq (4|\mathcal{X}|/r)^d$ balls of radius r > 0 [16, Prop. 5] with centers $\mathbf{z}_1, \ldots, \mathbf{z}_T \in \mathcal{X}_{\Delta}$. Since for all $\mathbf{z} \in \mathcal{X}_{\Delta}$ there exists $i \in \{1, \ldots, T\}$ such that $\|\mathbf{z} - \mathbf{z}_i\|_2 \leq r$, it holds for $f(\mathbf{z}) \coloneqq \partial^{\mathbf{p}, \mathbf{q}} \tilde{\mathbf{k}}(\mathbf{z}) - \partial^{\mathbf{p}, \mathbf{q}} \mathbf{k}(\mathbf{z})$ and $L_f \coloneqq \sup_{\mathbf{s} \in \mathcal{X}_{\Delta}} \|\nabla f(\mathbf{s})\|_2$ that

$$\left| f(\mathbf{z}) - f(\mathbf{z}_{i}) \right| \stackrel{\text{(a)}}{\leq} \sup_{\mathbf{s} \in \mathcal{X}_{\Delta}} \left\| \nabla f(\mathbf{s}) \right\|_{2} \left\| \mathbf{z} - \mathbf{z}_{i} \right\|_{2} \stackrel{\text{(b)}}{\leq} \sup_{\mathbf{s} \in \mathcal{X}_{\Delta}} \left\| \nabla f(\mathbf{s}) \right\|_{2} r = L_{f} r,$$
(31)

where (a) is due to the mean-value theorem followed by Cauchy-Schwarz inequality, and (b) is since $\|\mathbf{z} - \mathbf{z}_i\| \le r$. Now, by triangle inequality, it holds for any $\mathbf{z} \in \mathcal{X}_{\Delta}$ that

$$\left\|\nabla f(\mathbf{z})\right\|_{2} \leq \left\|\nabla \left[\partial^{\mathbf{p},\mathbf{q}}\tilde{\mathsf{k}}(\mathbf{z})\right]\right\|_{2} + \left\|\nabla \left[\partial^{\mathbf{p},\mathbf{q}}\mathsf{k}(\mathbf{z})\right]\right\|_{2} \leq \left\|\nabla \left[\partial^{\mathbf{p},\mathbf{q}}\tilde{\mathsf{k}}(\mathbf{z})\right]\right\|_{2} + D_{\mathbf{p},\mathbf{q},\mathcal{X}}.$$
(32)

Differentiating $\tilde{k}(z)$ as defined in (3), we get by the chain rule and triangle inequality that

$$\left\|\nabla\left[\partial^{\mathbf{p},\mathbf{q}}\tilde{\mathsf{k}}(\mathbf{z})\right]\right\|_{2} = \frac{1}{\tilde{d}} \left\|\sum_{j=1}^{\tilde{d}} \mathbf{w}_{j} \mathbf{w}_{j}^{\mathbf{p}+\mathbf{q}} \cos^{(|\mathbf{p}+\mathbf{q}|)}(\mathbf{w}_{j}^{\top}\mathbf{z})\right\|_{2} \le \frac{1}{\tilde{d}} \sum_{j=1}^{\tilde{d}} \left\|\mathbf{w}_{j}^{\mathbf{p}+\mathbf{q}}\right\| \left\|\mathbf{w}_{j}\right\|_{2}, \quad (33)$$

where $\cos^{(n)}$ denotes the *n*-th derivative of \cos for $n \in \mathbb{N}$. The main idea of the proof is then

$$\bigcap_{i=1}^{T} \{ |f(\mathbf{z}_i)| < \epsilon/2 \} \bigcap \{ L_f < \epsilon/2r \} \subseteq \{ |f(\mathbf{z})| < \epsilon : \forall \mathbf{z} \in \mathcal{X}_{\Delta} \},\$$

since $|f(\mathbf{z})| \leq |f(\mathbf{z}) - f(\mathbf{z}_i)| + |f(\mathbf{z}_i)|$ and (31) holds. Therefore, taking the complement and bounding the union, we get our governing inequality for the uniform error over \mathcal{X}_{Δ} :

$$\mathbb{P}\left[\sup_{\mathbf{z}\in\mathcal{X}_{\Delta}}\left|f(\mathbf{z})\right| \ge \epsilon\right] \le \sum_{i=1}^{T} \mathbb{P}\left[\left|f(\mathbf{z}_{i})\right| \ge \epsilon/2\right] + \mathbb{P}\left[L_{f} \ge \epsilon/2r\right].$$
(34)

Now, we need to bound all probabilities on the RHS. First, we deal with L_f ,

$$\mathbb{P}\left[L_{f} \geq \epsilon/2r\right]^{(c)} \leq \mathbb{E}\left[L_{f}\right] \frac{2r}{\epsilon} \leq \left(D_{\mathbf{p},\mathbf{q},\mathcal{X}} + \frac{1}{\tilde{d}} \sum_{j=1}^{\tilde{d}} \mathbb{E}_{\mathbf{w}_{j} \sim \Lambda}\left[\left\|\mathbf{w}_{j}^{\mathbf{p}+\mathbf{q}}\right\| \|\mathbf{w}_{j}\|_{2}\right]\right) \frac{2r}{\epsilon} \\ \leq \left(D_{\mathbf{p},\mathbf{q},\mathcal{X}} + E_{\mathbf{p},\mathbf{q}}\right) \frac{2r}{\epsilon},$$
(35)

where (c) is Markov's inequality, (d) is (32) and (33). To deal with the centers in (34), note $\partial^{\mathbf{p},\mathbf{q}}\tilde{k}(\mathbf{z})$ can be written as a sample average of \tilde{d} i.i.d. terms as per (3) since

$$\partial^{\mathbf{p},\mathbf{q}}\tilde{\mathbf{k}}(\mathbf{z}) = \partial^{\mathbf{p},\mathbf{q}}\left(\frac{1}{\tilde{d}}\sum_{i=1}^{\tilde{d}}\cos(\mathbf{w}_{j}^{\mathsf{T}}\mathbf{z})\right) = \frac{1}{\tilde{d}}\sum_{i=1}^{\tilde{d}}\partial^{\mathbf{p},\mathbf{q}}\cos(\mathbf{w}_{j}^{\mathsf{T}}\mathbf{z})$$

so that the Bernstein inequality (Cor. A.6) is applicable. For $j = 1, ..., \tilde{d}$, we have

$$\mathbb{E}_{\mathbf{w}_{j}\sim\Lambda}\left|\partial^{\mathbf{p},\mathbf{q}}\cos\left(\mathbf{w}_{j}^{\mathsf{T}}\mathbf{z}\right)\right|^{m} = \mathbb{E}_{\mathbf{w}_{j}\sim\Lambda}\left|\mathbf{w}_{j}^{\mathbf{p}+\mathbf{q}}\cos^{(|\mathbf{p}+\mathbf{q}|)}(\mathbf{w}_{j}^{\mathsf{T}}\mathbf{z})\right|^{m} \leq \mathbb{E}_{\mathbf{w}_{j}\sim\Lambda}\left|\mathbf{w}_{j}^{\mathbf{p}+\mathbf{q}}\right|^{m} \leq \frac{k!S^{2}R^{k-2}}{2},$$

and that $f(\mathbf{z}) = \partial^{\mathbf{p},\mathbf{q}}\tilde{\mathbf{k}}(\mathbf{z}) - \partial^{\mathbf{p},\mathbf{q}}\mathbf{k}(\mathbf{z}) = \partial^{\mathbf{p},\mathbf{q}}\tilde{\mathbf{k}}(\mathbf{z}) - \mathbb{E}[\partial^{\mathbf{p},\mathbf{q}}\tilde{\mathbf{k}}(\mathbf{z})]$ by the dominated convergence theorem. Hence, we may call the Bernstein inequality (Cor. A.6) to control $f(\mathbf{z}_i)$ so that

$$\mathbb{P}\left[\left|f(\mathbf{z}_{i})\right| \geq \epsilon/2\right] = \mathbb{P}\left[\left|\partial^{\mathbf{p},\mathbf{q}}\tilde{\mathsf{k}}(\mathbf{z}) - \partial^{\mathbf{p},\mathbf{q}}\mathsf{k}(\mathbf{z})\right| \geq \epsilon/2\right] \leq 2\exp\left(\frac{-\tilde{d}\epsilon^{2}}{4(2S^{2} + R\epsilon)}\right).$$
(36)

Combining the bounds for $|f(\mathbf{z}_i)|$ (36), and for L_f (35), into (34) yields

$$\mathbb{P}\left[\sup_{\mathbf{z}\in\mathcal{M}_{\Delta}}\left|f(\mathbf{z})\right|\geq\epsilon\right]\leq 2\left(\frac{4\left|\mathcal{X}\right|}{r}\right)^{d}\exp\left(\frac{-\tilde{d}\epsilon^{2}}{4(2S^{2}+R\epsilon)}\right)+\left(D_{\mathbf{p},\mathbf{q},\mathcal{X}}+E_{\mathbf{p},\mathbf{q}}\right)\frac{2r}{\epsilon},$$

which has the form $g(r) = \tau_1 r^{-d} + \tau_2 r$, that is minimized by choosing $r^* = (d\tau_1/\tau_2)^{\frac{1}{d+1}}$. This choice sets it to the form $g(r^*) = \tau_1^{\frac{1}{d+1}} \tau_2^{\frac{d}{d+1}} \left(d^{\frac{1}{d+1}} + d^{\frac{-d}{d+1}} \right)$, so that by substituting back in

$$\mathbb{P}\left[\sup_{\mathbf{z}\in\mathcal{M}_{\Delta}}\left|f(\mathbf{z})\right|\geq\epsilon\right]\leq F_{d}2^{\frac{3d+1}{d+1}}\left(\frac{|\mathcal{X}|(D_{\mathbf{p},\mathbf{q},\mathcal{X}}+E_{\mathbf{p},\mathbf{q}})}{\epsilon}\right)^{\frac{d}{d+1}}\exp\left(\frac{-\tilde{d}\epsilon^{2}}{4(d+1)(2S^{2}+R\epsilon)}\right).$$

Finally, we note that for $d \ge 1$, $F_d \coloneqq d^{\frac{1}{d+1}} + d^{\frac{-d}{d+1}} \le 2$ and $2^{\frac{3d+1}{d+1}} \le 8$.

C Bounds on Signature Kernels

We first set the ground for proving our main theorems by introducing the notion of *L*-Lipschitz kernels to help control distance distortions in the feature space. This subclass of kernels will be useful for us in relating the 1-variation of sequences in feature space to that in the input space. After this, we will prove various smaller lemmas and supplementary results for signature kernels, which will lead up to Lemma C.9, which is our main tool for proving Theorem 3.2, and it relates the concentration of our RFSF kernel \tilde{k}_{Sig} to the second derivatives of the RFF kernel \tilde{k} . Then, the proof of Theorem 3.2 quantifying the concentration of the RFSF kernel, \tilde{k}_{Sig} , will follow from putting together Lemma C.9 with Theorem B.1, and we will also make use of the Bernstein inequality from Theorem A.5. The proof of Theorem 3.5 for the RFSF-DP kernel, \tilde{k}_{Sig}^{DP} , will follow by combining the results of this section with α -exponential concentration, in particular, Theorem A.13. Finally, Theorem 3.8 for the RFSF-TRP kernel, \tilde{k}_{Sig}^{TRP} , will be proven using lemmas from this section, and the hypercontractivity concentration result from Theorem A.15.

Distance bounds in the RKHS.

Definition C.1 (Lipschitz kernel). *Let* (\mathfrak{X}, d) *be a metric space. We call a kernel* $\Bbbk : \mathfrak{X} \times \mathfrak{X} \to \mathbb{R}$ *with RKHS* \mathfrak{H} *an* L-*Lipschitz kernel over* \mathfrak{X} *for some* L > 0 *if it holds for all* $\mathbf{x}, \mathbf{y} \in \mathfrak{X}$ *that*

$$\left\|\mathbf{k}_{\mathbf{x}} - \mathbf{k}_{\mathbf{y}}\right\|_{\mathcal{H}} = \sqrt{\mathbf{k}(\mathbf{x}, \mathbf{x}) + \mathbf{k}(\mathbf{y}, \mathbf{y}) - 2\mathbf{k}(\mathbf{x}, \mathbf{y})} \le Ld(\mathbf{x}, \mathbf{y})$$

where $k_{\mathbf{x}} \coloneqq k(\mathbf{x}, \cdot), k_{\mathbf{y}} \coloneqq k(\mathbf{y}, \cdot) \in \mathcal{H}$.

Example C.2 (Finite 2nd spectral moment implies Lipschitz). Let $k : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ be a continuous, bounded and translation-invariant kernel with RKHS \mathcal{H} and spectral measure Λ , such that $\sigma_{\Lambda}^2 := \mathbb{E}_{\mathbf{w} \sim \Lambda} [\|\mathbf{w}\|_2^2] < \infty$. Then, it holds that k is $\|\mathbb{E}_{\mathbf{w} \sim \Lambda} [\mathbf{w} \mathbf{w}^{\top}]\|_2^{1/2}$ -Lipschitz, so for any $\mathbf{x}, \mathbf{y} \in \mathbb{R}^d$ one has

$$\|\mathbf{k}_{\mathbf{x}} - \mathbf{k}_{\mathbf{y}}\|_{\mathcal{H}} \leq \|\mathbb{E}_{\mathbf{w} \sim \Lambda}[\mathbf{w}\mathbf{w}^{\mathsf{T}}]\|_{2}^{1/2} \|\mathbf{x} - \mathbf{y}\|_{2}.$$

Proof. We have for $\mathbf{x}, \mathbf{y} \in \mathbb{R}^d$ that

$$\begin{aligned} \left\| \mathbf{k}_{\mathbf{x}} - \mathbf{k}_{\mathbf{y}} \right\|_{\mathcal{H}}^{2} \stackrel{(a)}{=} \mathbf{k}(\mathbf{x}, \mathbf{x}) + \mathbf{k}(\mathbf{y}, \mathbf{y}) - 2\mathbf{k}(\mathbf{x}, \mathbf{y}) \stackrel{(b)}{=} \int_{\mathbb{R}^{d}} 2 - 2\exp\left(i\mathbf{w}^{\top}(\mathbf{x} - \mathbf{y})\right) d\Lambda(\mathbf{w}) \\ \stackrel{(c)}{=} \int_{\mathbb{R}^{d}} 2 - 2\cos\left(\mathbf{w}^{\top}(\mathbf{x} - \mathbf{y})\right) d\Lambda(\mathbf{w}) \stackrel{(d)}{\leq} \int_{\mathbb{R}^{d}} \left(\mathbf{w}^{\top}(\mathbf{x} - \mathbf{y})\right)^{2} d\Lambda(\mathbf{w}) \\ \stackrel{(e)}{=} (\mathbf{x} - \mathbf{y})^{\top} \mathbb{E}_{\mathbf{w} \sim \Lambda} \left[\mathbf{w}\mathbf{w}^{\top}\right] (\mathbf{x} - \mathbf{y}) \stackrel{(f)}{\leq} \left\| \mathbb{E}_{\mathbf{w} \sim \Lambda} \left[\mathbf{w}\mathbf{w}^{\top}\right] \right\|_{2} \left\|\mathbf{x} - \mathbf{y}\right\|_{2}^{2}, \end{aligned}$$

where (a) holds due to the reproducing property, (b) is Bochner's theorem, (c) is because the imaginary part of the integral evaluates to 0 as the kernel is real-valued, (d) is due to the inequality $1 - t^2/2 \le \cos(t)$ for all $t \in \mathbb{R}$, (e) is because $(\mathbf{w}^{\top}(\mathbf{x}-\mathbf{y}))^2 = (\mathbf{x}-\mathbf{y})^{\top}\mathbf{w}\mathbf{w}^{\top}(\mathbf{x}-\mathbf{y})$, and (f) is Cauchy-Schwarz inequality combined with the definition of the spectral norm.

Example C.3 (Random Lipschitz bound for RFF). Let $\tilde{k} : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ be an RFF kernel defined as in (2) corresponding to some spectral measure Λ , and let $\tilde{\mathcal{H}}$ denote its feature space corresponding to the RFF map $\tilde{\varphi} : \mathbb{R}^d \to \tilde{\mathcal{H}}$ defined as in (2), so that given $\mathbf{w}_1, \ldots, \mathbf{w}_{\tilde{d}} \stackrel{i.i.d.}{\sim} \Lambda$, we have for $\mathbf{x}, \mathbf{y} \in \mathbb{R}^d$ that

$$\tilde{\mathsf{k}}(\mathbf{x},\mathbf{y}) = \frac{1}{\tilde{d}} \sum_{j=1}^{\tilde{d}} \cos\left(\mathbf{w}_{j}^{\mathsf{T}}(\mathbf{x}-\mathbf{y})\right).$$

Let $\mathbf{W} = (\mathbf{w}_1, \dots, \mathbf{w}_{\tilde{d}}) \in \mathbb{R}^{d \times \tilde{d}}$ be the random matrix with column vectors $\mathbf{w}_1, \dots, \mathbf{w}_{\tilde{d}} \stackrel{i.i.d.}{\sim} \Lambda$. Then, \tilde{k} is $\left(\frac{\|\mathbf{W}\|_2}{\sqrt{\tilde{d}}}\right)$ -Lipschitz, so that for any $\mathbf{x}, \mathbf{y} \in \mathbb{R}^d$, we have the inequality

$$\left\| \tilde{\varphi}(\mathbf{x}) - \tilde{\varphi}(\mathbf{y}) \right\|_{\tilde{\mathcal{H}}} \leq \frac{\|\mathbf{W}\|_2}{\sqrt{\tilde{d}}} \left\| \mathbf{x} - \mathbf{y} \right\|_2.$$

Proof. The proof follows analogously to that of Example C.2. Let $\mathbf{x}, \mathbf{y} \in \mathbb{R}^d$, then

$$\begin{aligned} \left\| \tilde{\varphi}(\mathbf{x}) - \tilde{\varphi}(\mathbf{y}) \right\|_{\tilde{\mathcal{H}}}^{2} &\stackrel{\text{(a)}}{=} 2 - \frac{2}{\tilde{d}} \sum_{i=1}^{\tilde{d}} \cos(\mathbf{w}_{i}^{\top}(\mathbf{x} - \mathbf{y})) \stackrel{\text{(b)}}{\leq} \frac{1}{\tilde{d}} \sum_{i=1}^{\tilde{d}} \left(\mathbf{w}_{i}^{\top}(\mathbf{x} - \mathbf{y}) \right)^{2} \\ &\stackrel{\text{(c)}}{=} \frac{1}{\tilde{d}} (\mathbf{x} - \mathbf{y})^{\top} \mathbf{W} \mathbf{W}^{\top}(\mathbf{x} - \mathbf{y}) \stackrel{\text{(d)}}{\leq} \frac{1}{\tilde{d}} \left\| \mathbf{W} \right\|_{2}^{2} \left\| \mathbf{x} - \mathbf{y} \right\|_{2}^{2}, \end{aligned}$$

where (a) is due to the cosine identity $\cos(a-b) = \cos(a)\cos(b) + \sin(a)\sin(b)$ for all $a, b \in \mathbb{R}$, (b) is due to the inequality $1 - t^2/2 \le \cos(t)$ for all $t \in \mathbb{R}$, (c) is because $(\mathbf{w}_i^{\mathsf{T}}(\mathbf{x}-\mathbf{y}))^2 = (\mathbf{x}-\mathbf{y})^{\mathsf{T}}\mathbf{w}_i\mathbf{w}_i^{\mathsf{T}}(\mathbf{x}-\mathbf{y})$ and $\mathbf{W}\mathbf{W}^{\mathsf{T}} = \sum_{i=1}^{\tilde{d}} \mathbf{w}_i\mathbf{w}_i^{\mathsf{T}}$, (d) is due to the Cauchy-Schwarz inequality combined with the definition of the spectral norm.

Bounds for the Signature Kernel. A well-known property of signature features that they decay factorially fast with respect to the tensor level $m \in \mathbb{N}$.

Lemma C.4 (Norm bound for signature features). Let L > 0 and $k : \mathfrak{X} \times \mathfrak{X} \to \mathbb{R}$ be an *L*-Lipschitz kernel with RKHS \mathfrak{H} . Then, we have for the level-m signature features $\varphi_{\text{Sig}_m}(\mathbf{x})$ of the sequence $\mathbf{x} \in \mathfrak{X}_{\text{seq}}$ that

$$\left\|\varphi_{\operatorname{Sig}_m}(\mathbf{x})\right\|_{\mathcal{H}^{\otimes m}} \leq \frac{\left(L \|\mathbf{x}\|_{1-\operatorname{var}}\right)^m}{m!}.$$

Proof. We have

$$\begin{aligned} \left\| \varphi_{\mathrm{Sig}_{m}}(\mathbf{x}) \right\|_{\mathcal{H}^{\otimes m}} &= \left\| \sum_{\mathbf{i} \in \Delta_{m}(\ell_{\mathbf{x}}-1)} \delta \mathbf{k}_{\mathbf{x}_{i_{1}}} \otimes \cdots \otimes \delta \mathbf{k}_{\mathbf{x}_{i_{m}}} \right\|_{\mathcal{H}^{\otimes m}} \\ &\stackrel{(a)}{\leq} \sum_{\mathbf{i} \in \Delta_{m}(\ell_{\mathbf{x}}-1)} \left\| \delta \mathbf{k}_{\mathbf{x}_{i_{1}}} \right\|_{\mathcal{H}} \cdots \left\| \delta \mathbf{k}_{\mathbf{x}_{i_{m}}} \right\|_{\mathcal{H}} \\ &\stackrel{(b)}{\leq} L^{m} \sum_{\mathbf{i} \in \Delta_{m}(|\mathbf{x}-1|)} \left\| \delta \mathbf{x}_{i_{1}} \right\|_{2} \cdots \left\| \delta \mathbf{x}_{i_{m}} \right\|_{2} \\ &\stackrel{(c)}{\leq} \frac{\left(L \sum_{i=1}^{\ell_{\mathbf{x}}-1} \left\| \delta \mathbf{x}_{i} \right\|_{2} \right)^{m}}{m!} \stackrel{(d)}{=} \frac{(L \|\mathbf{x}\|_{1-\mathrm{var}})^{m}}{m!} \end{aligned}$$

where (a) follows from triangle inequality followed by factorizing the tensor norm, (b) is the *L*-Lipschitzness property, (c) from completing the multinomial expansion and normalizing by the number of permutations (note that $\Delta_m(\ell_x - 1)$ contains a single permutation of all such multi-indices that have nonrepeating entries), and (d) is the definition of sequence 1-variation.

The following bound for the signature kernel is a direct consequence of the previous lemma.

Corollary C.5 (Upper bound for signature kernel). *Let* $k: \mathfrak{X} \times \mathfrak{X} \to \mathbb{R}$ *be an L*-*Lipschitz kernel, and* $k_{\text{Sig}_m}: \mathfrak{X}_{\text{seq}} \times \mathfrak{X}_{\text{seq}} \to \mathbb{R}$ *the level-m signature kernel built from* k. *Then, we have the following bound for* $\mathbf{x}, \mathbf{y} \in \mathfrak{X}_{\text{seq}}$

$$\left|\mathsf{k}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y})\right| \leq \frac{\left(L^{2} \|\mathbf{x}\|_{1-\mathrm{var}} \|\mathbf{y}\|_{1-\mathrm{var}}\right)^{m}}{(m!)^{2}}.$$

Proof. Note that without a kernel trick, k_{Sig_m} is written for $\mathbf{x}, \mathbf{y} \in \mathcal{X}_{\text{seq}}$ as the inner product

$$\begin{split} \left| \mathsf{k}_{\mathrm{Sig}_{m}}(\mathbf{x}, \mathbf{y}) \right| &= \left| \left\langle \sum_{\mathbf{i} \in \Delta_{m}(\ell_{\mathbf{x}}-1)} \delta \mathsf{k}_{\mathbf{x}_{i_{1}}} \otimes \cdots \otimes \delta \mathsf{k}_{\mathbf{x}_{i_{m}}}, \sum_{\mathbf{j} \in \Delta_{m}(\ell_{\mathbf{y}}-1)} \delta \mathsf{k}_{\mathbf{y}_{j_{1}}} \otimes \cdots \otimes \delta \mathsf{k}_{\mathbf{y}_{j_{m}}} \right\rangle_{\mathcal{H}^{\otimes m}} \right| \\ & \stackrel{(a)}{\leq} \left\| \sum_{\mathbf{i} \in \Delta_{m}(\ell_{\mathbf{x}}-1)} \delta \mathsf{k}_{\mathbf{x}_{i_{1}}} \otimes \cdots \otimes \delta \mathsf{k}_{\mathbf{x}_{i_{m}}} \right\|_{\mathcal{H}^{\otimes m}} \left\| \sum_{\mathbf{j} \in \Delta_{m}(\ell_{\mathbf{y}}-1)} \delta \mathsf{k}_{\mathbf{y}_{j_{1}}} \otimes \cdots \otimes \delta \mathsf{k}_{\mathbf{y}_{j_{m}}} \right\|_{\mathcal{H}^{\otimes m}} \\ & \stackrel{(b)}{\leq} \frac{\left(L^{2} \|\mathbf{x}\|_{1-\mathrm{var}} \|\mathbf{y}\|_{1-\mathrm{var}}\right)^{m}}{(m!)^{2}}, \end{split}$$

where (a) follows from the Cauchy-Schwarz inequality, and (b) is implied by Lemma C.4. \Box

A similar upper bound to Lemma C.4 also holds for the RFSF kernel \tilde{k}_{Sig_m} , that now depends on the norms of the random matrices $\mathbf{W}^{(1)}, \ldots, \mathbf{W}^{(m)}$, hence is itself random.

Lemma C.6 (Random norm bound for RFSF). Let $\tilde{\varphi}_{\text{Sig}_m} : \mathfrak{X}_{\text{seq}} \to \tilde{\mathcal{H}}_{\text{Sig}}$ be the level-*m* RFSF map defined as in (13) built from some spectral measure Λ . Then, we have for $\mathbf{x} \in \mathfrak{X}_{\text{seq}}$ that

$$\left\| \tilde{\varphi}_{\operatorname{Sig}_{m}}(\mathbf{x}) \right\|_{\tilde{\mathcal{H}}(\otimes m)} \leq \frac{\left\| \mathbf{W}^{(1)} \right\|_{2} \cdots \left\| \mathbf{W}^{(m)} \right\|_{2}}{m!} \left(\frac{\left\| \mathbf{x} \right\|_{1-\operatorname{var}}}{\sqrt{\tilde{d}}} \right)^{m},$$

where $\mathbf{W}^{(1)}, \dots, \mathbf{W}^{(m)} \stackrel{i.i.d.}{\sim} \Lambda^{\tilde{d}}$ are random matrices sampled from $\Lambda^{\tilde{d}}$.

Proof. The proof follows analogously to Lemma C.4. We have for $\mathbf{x} \in \mathcal{X}_{seq}$ that

$$\begin{split} \left\| \tilde{\varphi}_{\mathrm{Sig}_{m}}(\mathbf{x}) \right\|_{\tilde{\mathcal{H}}(\otimes m} &= \left\| \sum_{\mathbf{i} \in \Delta_{m}(\ell_{\mathbf{x}}-1)} \delta \tilde{\varphi}_{1}(\mathbf{x}_{i_{1}}) \otimes \cdots \otimes \delta \tilde{\varphi}_{m}(\mathbf{x}_{i_{m}}) \right\|_{\tilde{\mathcal{H}}(\otimes m)} \\ &\stackrel{(a)}{\leq} \sum_{\mathbf{i} \in \Delta_{m}(\ell_{\mathbf{x}}-1)} \left\| \delta \tilde{\varphi}_{1}(\mathbf{x}_{i_{1}}) \right\|_{\tilde{\mathcal{H}}} \cdots \left\| \delta \tilde{\varphi}_{m}(\mathbf{x}_{i_{m}}) \right\|_{\tilde{\mathcal{H}}} \\ &\stackrel{(b)}{\leq} \frac{\left\| \mathbf{W}^{(1)} \right\|_{2} \cdots \left\| \mathbf{W}^{(m)} \right\|_{2}}{\tilde{d}^{m/2}} \sum_{\mathbf{i} \in \Delta_{m}(\ell_{\mathbf{x}}-1)} \left\| \delta \mathbf{x}_{i_{1}} \right\|_{2} \cdots \left\| \delta \mathbf{x}_{i_{m}} \right\|_{2} \\ &\stackrel{(c)}{\leq} \frac{\left\| \mathbf{W}^{(1)} \right\|_{2} \cdots \left\| \mathbf{W}^{(m)} \right\|_{2}}{\tilde{d}^{m/2}} \frac{\| \mathbf{x} \|_{1\text{-var}}^{m}}{m!}, \end{split}$$

where (a) is the triangle inequality and factorization of tensor norm, (b) is using the Lipschitzness of RFFs from Example C.3, (c) is the same as steps (c)-(d) in Lemma C.4.

Then, the following is again an application of the Cauchy-Schwarz inequality.

Corollary C.7 (Random upper bound for RFSF kernel). Let $\tilde{k}_{Sig_m} : \mathfrak{X}_{seq} \times \mathfrak{X}_{seq} \to \mathbb{R}$ be the level-*m* RFSF kernel defined as in (14) built from some spectral measure Λ . Then, for all $\mathbf{x}, \mathbf{y} \in \mathfrak{X}_{seq}$

$$\left|\tilde{\mathsf{K}}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y})\right| \leq \frac{\left\|\mathbf{W}^{(1)}\right\|_{2}^{2} \cdots \left\|\mathbf{W}^{(m)}\right\|_{2}^{2}}{(m!)^{2}} \left(\frac{\left\|\mathbf{x}\right\|_{1} \cdot \operatorname{var}}{\tilde{d}} \left\|\mathbf{y}\right\|_{1} \cdot \operatorname{var}}{\tilde{d}}\right)^{m}$$

where $\mathbf{W}^{(1)}, \ldots, \mathbf{W}^{(m)} \stackrel{i.i.d.}{\sim} \Lambda^{\tilde{d}}$ are random matrices sampled from $\Lambda^{\tilde{d}}$.

The following lemma does not concern signatures, but will be useful to us later in the proof of Lemma C.9 by providing a mean-value theorem for the cross-differencing operator $\delta_{i,j}^2$.

Lemma C.8 (2nd order mean-value theorem). Let $f : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ be a twice differentiable function, and $\mathfrak{X} \subset \mathbb{R}^d$ be a convex and compact set. Then, we have for any $\mathbf{u}, \mathbf{v}, \mathbf{x}, \mathbf{y} \in \mathfrak{X}$ that

$$f(\mathbf{x},\mathbf{y}) - f(\mathbf{x},\mathbf{v}) - f(\mathbf{u},\mathbf{y}) + f(\mathbf{u},\mathbf{v}) \leq \sup_{\mathbf{s},\mathbf{t}\in\mathcal{X}} \left\| \partial_{\mathbf{s},\mathbf{t}}^2 f(\mathbf{s},\mathbf{t}) \right\|_2 \|\mathbf{x}-\mathbf{u}\|_2 \|\mathbf{y}-\mathbf{v}\|_2,$$

where $\partial_{\mathbf{s},\mathbf{t}}^2 f(\mathbf{s},\mathbf{t}) \coloneqq \left(\frac{\partial^2 f(\mathbf{s},\mathbf{t})}{\partial s_i \partial t_j}\right)_{i,j=1}^d$ refers to the submatrix of the Hessian of cross-derivatives.

Proof. Keeping $\mathbf{v}, \mathbf{y} \in \mathcal{X}$ as fixed, we may define $g : \mathbb{R}^d \to \mathbb{R}$ as $g(\cdot) \coloneqq f(\cdot, \mathbf{y}) - f(\cdot, \mathbf{v})$, so that the expression above can be written as $g(\mathbf{x}) - g(\mathbf{u})$, and by the convexity of \mathcal{X} , we may apply the mean-value theorem to find that $\exists s \in (0, 1)$ such that

$$g(\mathbf{x}) - g(\mathbf{u}) = \left\langle \nabla g(s\mathbf{x} + (1 - s)\mathbf{u}), \mathbf{x} - \mathbf{u} \right\rangle \le \sup_{\mathbf{s} \in \mathcal{X}} \left\| \nabla g(\mathbf{s}) \right\|_2 \|\mathbf{x} - \mathbf{u}\|_2,$$
(37)

where $\nabla g(\mathbf{s}) \coloneqq \left(\frac{\partial g(\mathbf{s})}{\partial s_i}\right)_{i=1}^d$ denotes the gradient of g, while the second inequality follows from the Cauchy-Schwarz inequality and the compactness of \mathcal{X} (so that the sup exists). Also note that $\nabla g(\mathbf{s}) = \partial_{\mathbf{s}} f(\mathbf{s}, \mathbf{y}) - \partial_{\mathbf{s}} f(\mathbf{s}, \mathbf{v})$, so defining $h : \mathbb{R}^d \to \mathbb{R}^d$ as $h(\cdot) \coloneqq \partial_{\mathbf{s}} f(\mathbf{s}, \cdot)$ and applying the vector-valued mean-value inequality [66, Thm. 9.19] to h gives that $\exists t \in (0, 1)$ such that

$$\left\|h(\mathbf{y}) - h(\mathbf{v})\right\|_{2} \le \left\|\mathbf{J}_{h}(t\mathbf{y} + (1-t)\mathbf{v})\right\|_{2} \left\|\mathbf{y} - \mathbf{v}\right\|_{2} \le \sup_{\mathbf{t}\in\mathcal{X}} \left\|\mathbf{J}_{h}(\mathbf{t})\right\|_{2} \left\|\mathbf{y} - \mathbf{v}\right\|_{2},$$
(38)

where $J_h(\mathbf{t}) \coloneqq \left(\frac{\partial h_i(\mathbf{t})}{\partial t_j}\right)_{i,j=1}^d$ refers to the Jacobian of *h*. Putting the inequalities (37) and (38) together and substituting back the function *f* gives the desired result.

This is our final lemma in our exposition of supplementary results about the (random) signature kernels, and it will be our main tool for proving Theorem 3.2.

Lemma C.9 (Uniform upper bound for deviation of RFSF kernel). Let $\mathfrak{X} \subset \mathbb{R}^d$ be a convex and compact set, $k : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ a continuous, bounded, translation-invariant *L*-Lipschitz kernel and $\tilde{k} : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ the corresponding RFF kernel. Then, the level-m ($m \in \mathbb{N}$) signature and RFSF kernels are uniformly close for V > 0 by

$$\sup_{\substack{\mathbf{x},\mathbf{y}\in\mathcal{X}_{\text{seq}}\\\|\|\mathbf{x}\|_{1-\text{var}},\|\|\mathbf{y}\|_{1-\text{var}}\leq V}} \left\| \mathbf{w}^{(1)} \|_{2}^{2} \cdots \|\mathbf{w}^{(k-1)}\|_{2}^{2} \sup_{\mathbf{s},\mathbf{t}\in\mathcal{X}} \|\partial_{\mathbf{s},\mathbf{t}}^{2}\tilde{\mathbf{k}}_{k}(\mathbf{s},\mathbf{t}) - \partial_{\mathbf{s},\mathbf{t}}^{2}\mathbf{k}(\mathbf{s},\mathbf{t})\|_{2}^{2},$$

where $\tilde{k}_1, \ldots, \tilde{k}_m$ are independent RFF kernels with weights $\mathbf{W}^{(1)}, \ldots, \mathbf{W}^{(m)} \stackrel{i.i.d.}{\sim} \Lambda^{\tilde{d}}$, and $\partial_{\mathbf{s},\mathbf{t}}^2 f(\mathbf{s},\mathbf{t}) := \left(\frac{\partial^2 f(\mathbf{s},\mathbf{t})}{\partial s_i \partial t_j}\right)_{i,j=1}^d$ for a twice-differentiable function $f : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$.

Proof. First of all, by Lemma C.5 and Lemma C.7, the supremum exists. In the following, given a sequence $\mathbf{x} \in \mathcal{X}_{seq}$, we denote its 1 : l slice for some $l \in [l_{\mathbf{x}}] \mathbf{x}_{1:l} \coloneqq (\mathbf{x}_1, \dots, \mathbf{x}_l)$. Then, it holds for any $m \ge 1$ recursively for the signature kernel that

$$k_{\text{Sig}_{m}}(\mathbf{x}, \mathbf{y}) = \sum_{k=1}^{\ell_{x}-1} \sum_{l=1}^{\ell_{y}-1} k_{\text{Sig}_{m-1}}(\mathbf{x}_{1:k}, \mathbf{y}_{1:l}) \delta_{k,l}^{2} k(\mathbf{x}_{k}, \mathbf{y}_{l}),$$

and analogously for the RFSF kernel that

$$\tilde{\mathsf{k}}_{\mathrm{Sig}_m}(\mathbf{x},\mathbf{y}) = \sum_{k=1}^{\ell_{\mathbf{x}}-1} \sum_{l=1}^{\ell_{\mathbf{y}}-1} \tilde{\mathsf{k}}_{\mathrm{Sig}_{m-1}}(\mathbf{x}_{1:k},\mathbf{y}_{1:l}) \delta_{k,l}^2 \tilde{\mathsf{k}}_m(\mathbf{x}_k,\mathbf{y}_l).$$

Combining these recursions together, we have for the uniform error that

$$\begin{split} \epsilon_{m} &= \sup_{\substack{\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq} \\ \|\mathbf{x}\|_{1-var}, \|\mathbf{y}\|_{1-var} \leq V}} \left| \tilde{k}_{Sig_{m}}(\mathbf{x}, \mathbf{y}) - k_{Sig_{m}}(\mathbf{x}, \mathbf{y}) \right| \\ &= \sup_{\substack{\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq} \\ \|\mathbf{x}\|_{1-var}, \|\mathbf{y}\|_{1-var} \leq V}} \left| \sum_{k=1}^{\ell_{x}-1} \sum_{l=1}^{\ell_{y}-1} \tilde{k}_{Sig_{m-1}}(\mathbf{x}_{1:k}, \mathbf{y}_{1:l}) \delta_{k,l}^{2} \tilde{k}_{m}(\mathbf{x}_{k}, \mathbf{y}_{l}) - k_{Sig_{m-1}}(\mathbf{x}_{1:k}, \mathbf{y}_{1:l}) \delta_{k,l}^{2} k(\mathbf{x}_{k}, \mathbf{y}_{l}) \right| \\ &\stackrel{(a)}{\leq} \sup_{\substack{\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq} \\ \|\mathbf{x}\|_{1-var}, \|\mathbf{y}\|_{1-var} \leq V}} \left| \sum_{k=1}^{\ell_{x}-1} \sum_{l=1}^{\ell_{y}-1} \tilde{k}_{Sig_{m-1}}(\mathbf{x}_{1:k}, \mathbf{y}_{1:l}) (\delta_{k,l}^{2} \tilde{k}_{m}(\mathbf{x}_{k}, \mathbf{y}_{l}) - \delta_{k,l}^{2} k(\mathbf{x}_{k}, \mathbf{y}_{l})) \right| \\ &+ \sup_{\substack{\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq} \\ \|\mathbf{x}\|_{1-var}, \|\mathbf{y}\|_{1-var} \leq V}} \left| \sum_{k=1}^{\ell_{x}-1} \sum_{l=1}^{\ell_{y}-1} (\tilde{k}_{Sig_{m-1}}(\mathbf{x}_{1:k}, \mathbf{y}_{1:l}) - \tilde{k}_{Sig_{m-1}}(\mathbf{x}_{1:k}, \mathbf{y}_{1:l})) \delta_{k,l}^{2} k(\mathbf{x}_{k}, \mathbf{y}_{l}) \right| \end{aligned}$$

$$\stackrel{\text{(b)}}{\leq} \sup_{\substack{\mathbf{x},\mathbf{y}\in\mathcal{X}_{\text{seq}}\\\|\|\mathbf{x}\|_{1-\text{var}}\|\|\mathbf{y}\|_{1-\text{var}}\leq V}} \sum_{k=1}^{\ell_{x}-1} \sum_{l=1}^{\ell_{y}-1} \left| \tilde{k}_{\text{Sig}_{m-1}}(\mathbf{x}_{1:k},\mathbf{y}_{1:l}) \right| \left| \delta_{k,l}^{2} \tilde{k}_{m}(\mathbf{x}_{k},\mathbf{y}_{l}) - \delta_{k,l}^{2} \mathbf{k}(\mathbf{x}_{k},\mathbf{y}_{l}) \right|$$

$$+ \sup_{\substack{\mathbf{x},\mathbf{y}\in\mathcal{X}_{\text{seq}}\\\|\|\mathbf{x}\|_{1-\text{var}}\|\|\mathbf{y}\|_{1-\text{var}}\leq V \\ (\|\mathbf{x}\|_{1-\text{var}},\|\|\mathbf{y}\|_{1-\text{var}}\leq V } \sum_{k=1}^{\ell_{x}-1} \sum_{l=1}^{\ell_{y}-1} \left| \tilde{k}_{\text{Sig}_{m-1}}(\mathbf{x}_{1:k},\mathbf{y}_{1:l}) - \mathbf{k}_{\text{Sig}_{m-1}}(\mathbf{x}_{1:k},\mathbf{y}_{1:l}) \right| \left| \delta_{k,l}^{2} \mathbf{k}(\mathbf{x}_{k},\mathbf{y}_{l}) \right|$$

$$+ \sup_{\substack{\mathbf{x},\mathbf{y}\in\mathcal{X}_{\text{seq}}\\\|\|\mathbf{x}\|_{1-\text{var}},\|\|\mathbf{y}\|_{1-\text{var}}\leq V \\ (\|\mathbf{x}\|_{1-\text{var}},\|\|\mathbf{y}\|_{1-\text{var}}\leq V } \left| \tilde{k}_{\text{Sig}_{m-1}}(\mathbf{x},\mathbf{y}) - \mathbf{k}_{\text{Sig}_{m-1}}(\mathbf{x},\mathbf{y}) \right| \underbrace{\sup_{\mathbf{x},\mathbf{y}\in\mathcal{X}_{\text{seq}}}_{(\|\mathbf{x}\|_{1-\text{var}},\|\mathbf{y}\|_{1-\text{var}}\leq V \\ (\|\mathbf{x}\|_{1-\text{var}},\|\mathbf{y}\|_{1-\text{var}}\leq V \\ (\|\mathbf{x}\|_{1-\text{var}},\|\mathbf{y}\|_{1-\text{var}}\leq V \\ (\|\mathbf{x}\|_{1-\text{var}},\|\mathbf{y}\|_{1-\text{var}}\leq V \\ (\|\mathbf{x}\|_{1-\text{var},\|\mathbf{y}\|_{1-\text{var}}\leq V \\ (\|\mathbf{x}\|_{1-\text{var},\|\mathbf{x}\|_{1-\text{var},\|\mathbf{y}\|_{1-\text{var}}\leq V \\ (\|\mathbf{x}\|_{1-\text{var},\|\mathbf{x}\|_{1-\text{var}}$$

where (a) follows from adding and subtracting the cross-terms and applying triangle inequality, (b) follows from applying triangle inequality over the summations, (c) follows from noting that if $\|\mathbf{x}\|_{1-\text{var}}$, $\|\mathbf{y}\|_{1-\text{var}} \leq V$ then so is $\|\mathbf{x}_{1:k}\|_{1-\text{var}}$, $\|\mathbf{y}\|_{1-\text{var}} \leq V$ for $k \in [\ell_x]$ and $l \in [\ell_y]$, and thus justifiably pulling out the supremums.

Now, we deal with terms (i)-(iv) individually. For (i), we have Corollary C.7, so

$$\sup_{\substack{\mathbf{x},\mathbf{y}\in\mathcal{X}_{\text{seq}}\\\|\mathbf{x}\|_{1\text{-var}},\|\mathbf{y}\|_{1\text{-var}}\leq V}} \left|\tilde{k}_{\text{Sig}_{m-1}}(\mathbf{x},\mathbf{y})\right| \leq \frac{\left\|\mathbf{W}^{(1)}\right\|_{2}^{2}\cdots\left\|\mathbf{W}^{(m-1)}\right\|_{2}^{2}}{((m-1)!)^{2}} \left(\frac{V^{2}}{\tilde{d}}\right)^{m-1}.$$
(40)

To deal with (ii), we can apply Lemma C.8 with $f = \tilde{k}_m - k$ to get

$$\sup_{\substack{\mathbf{x},\mathbf{y}\in\mathcal{X}_{\mathrm{seq}}\\\|\mathbf{x}_{\|_{1-\mathrm{var}}}\|\mathbf{y}\|_{1-\mathrm{var}}\leq V}} \sum_{k=1}^{\ell_{\mathbf{x}}-1} \sum_{l=1}^{\ell_{\mathbf{y}}-1} \left| \delta_{k,l}^{2} \tilde{\mathbf{k}}_{m}(\mathbf{x}_{k},\mathbf{y}_{l}) - \delta_{k,l}^{2} \mathbf{k}(\mathbf{x}_{k},\mathbf{y}_{l}) \right| \\
\leq \sup_{\mathbf{s},\mathbf{t}\in\mathcal{X}} \left\| \partial_{\mathbf{s},\mathbf{t}}^{2} \tilde{\mathbf{k}}_{m}(\mathbf{s},\mathbf{t}) - \partial_{\mathbf{s},\mathbf{t}}^{2} \mathbf{k}(\mathbf{s},\mathbf{t}) \right\|_{2} \sup_{\substack{\mathbf{x},\mathbf{y}\in\mathcal{X}_{\mathrm{seq}}\\\|\mathbf{x}\|_{1-\mathrm{var}},\|\mathbf{y}\|_{1-\mathrm{var}}\leq V}} \sum_{k=1}^{\ell_{\mathbf{x}}-1} \sum_{l=1}^{\ell_{\mathbf{y}}-1} \|\delta\mathbf{x}_{k}\|_{2} \left\| \delta\mathbf{y}_{l} \right\|_{2} \\
\leq V^{2} \sup_{\mathbf{s},\mathbf{t}\in\mathcal{X}} \left\| \partial_{\mathbf{s},\mathbf{t}}^{2} \tilde{\mathbf{k}}_{m}(\mathbf{s},\mathbf{t}) - \partial_{\mathbf{s},\mathbf{t}}^{2} \mathbf{k}(\mathbf{s},\mathbf{t}) \right\|_{2}.$$
(41)

For (iii), we note that it is simply ϵ_{m-1} . Finally, we can write (iv) as an inner product and apply Cauchy-Schwarz and *L*-Lipschitzness of k so that

$$\sup_{\substack{\mathbf{x}, \mathbf{y} \in \mathcal{X}_{\text{seq}} \\ \|\mathbf{x}\|_{1-\text{var}}, \|\mathbf{y}\|_{1-\text{var}} \leq V}} \sum_{k=1}^{\ell_{\mathbf{x}}-1} \sum_{l=1}^{\ell_{\mathbf{y}}-1} \left| \delta_{k,l}^{2} \mathsf{k}(\mathbf{x}_{k}, \mathbf{y}_{l}) \right| = \sup_{\substack{\mathbf{x}, \mathbf{y} \in \mathcal{X}_{\text{seq}} \\ \|\mathbf{x}\|_{1-\text{var}}, \|\mathbf{y}\|_{1-\text{var}} \leq V}} \sum_{k=1}^{\ell_{\mathbf{x}}-1} \sum_{l=1}^{\ell_{\mathbf{y}}-1} \left\| \delta_{k,l}^{2} \mathsf{k}(\mathbf{x}_{k}, \mathbf{y}_{l}) \right\|_{1-\text{var}}} \left\| \mathbf{y} \right\|_{1-\text{var}} \leq V \sum_{k=1}^{\ell_{\mathbf{x}}-1} \sum_{l=1}^{\ell_{\mathbf{y}}-1} \left\| \delta_{k,k} \right\|_{\mathcal{H}} \left\| \delta_{k,k} \right\|_{\mathcal{H}} \left\| \delta_{k,k} \right\|_{\mathcal{H}} \leq L^{2} V^{2}.$$

$$(42)$$

Putting equations (40), (41), (42) together in (39), we get that

$$\epsilon_{m} = \sup_{\substack{\mathbf{x}, \mathbf{y} \in \mathcal{X}_{\text{seq}} \\ \|\mathbf{x}\|_{1-\text{var}}, \|\mathbf{y}\|_{1-\text{var}} \leq V}} \left| \tilde{k}_{\text{Sig}_{m}}(\mathbf{x}, \mathbf{y}) - k_{\text{Sig}_{m}}(\mathbf{x}, \mathbf{y}) \right|$$

$$\leq \frac{V^{2m}}{\tilde{d}^{m-1}((m-1)!)^{2}} \sup_{\mathbf{s}, \mathbf{t} \in \mathcal{X}} \left\| \partial_{\mathbf{s}, \mathbf{t}}^{2} \tilde{k}_{m}(\mathbf{s}, \mathbf{t}) - \partial_{\mathbf{s}, \mathbf{t}}^{2} k(\mathbf{s}, \mathbf{t}) \right\|_{2} \prod_{p=1}^{m-1} \left\| \mathbf{W}^{(p)} \right\|_{2}^{2} + L^{2} V^{2} \epsilon_{m-1}, \quad (43)$$

which gives us a recursion for estimating ϵ_m . The initial step, m = 1, can be estimated by

$$\begin{aligned} \epsilon_{1} &= \sup_{\substack{\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq} \\ \|\mathbf{x}\|_{1-\operatorname{var}}, \|\mathbf{y}\|_{1-\operatorname{var}} \leq V}} \left| \tilde{k}_{\operatorname{Sig}_{1}}(\mathbf{x}, \mathbf{y}) - k_{\operatorname{Sig}_{1}}(\mathbf{x}, \mathbf{y}) \right| \\ &= \sup_{\substack{\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq} \\ \|\mathbf{x}\|_{1-\operatorname{var}}, \|\mathbf{y}\|_{1-\operatorname{var}} \leq V}} \left| \sum_{k=1}^{\ell_{\mathbf{x}}-1} \sum_{l=1}^{\ell_{\mathbf{y}}-1} \delta_{k,l}^{2} \tilde{k}_{1}(\mathbf{x}_{k}, \mathbf{y}_{l}) - \delta_{k,l}^{2} \mathbf{k}(\mathbf{x}_{k}, \mathbf{y}_{l}) \right| \\ &= \sup_{\substack{\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq} \\ \|\mathbf{x}\|_{1-\operatorname{var}}, \|\mathbf{y}\|_{1-\operatorname{var}} \leq V}} \sum_{k=1}^{\ell_{\mathbf{x}}-1} \sum_{l=0}^{\ell_{\mathbf{y}}-1} \left| \delta_{k,l}^{2} \tilde{k}_{1}(\mathbf{x}_{k}, \mathbf{y}_{l}) - \delta_{k,l}^{2} \mathbf{k}(\mathbf{x}_{k}, \mathbf{y}_{l}) \right| \\ &\leq V^{2} \sup_{\mathbf{s}, \mathbf{t} \in \mathcal{X}} \left\| \partial_{\mathbf{s},\mathbf{t}}^{2} \tilde{k}_{1}(\mathbf{s}, \mathbf{t}) - \partial_{\mathbf{s},\mathbf{t}}^{2} \mathbf{k}(\mathbf{s}, \mathbf{t}) \right\|_{2}, \end{aligned}$$
(44)

which is actually analogous to (43) since $\epsilon_0 = 0$. Now, we may unroll the recursion (43) with the initial condition (44), and we get

$$\epsilon_m \leq V^{2m} \sum_{k=1}^m \frac{L^{2(m-k)}}{\tilde{d}^{k-1}((k-1)!)^2} \sup_{\mathbf{s},\mathbf{t}\in\mathcal{X}} \left\|\partial_{\mathbf{s},\mathbf{t}}^2 \tilde{k}_k(\mathbf{s},\mathbf{t}) - \partial_{\mathbf{s},\mathbf{t}}^2 \mathbf{k}(\mathbf{s},\mathbf{t})\right\|_2 \prod_{p=1}^{k-1} \left\|\mathbf{W}^{(p)}\right\|_2^2.$$

Proofs of main concentration results. Here, we provide proofs of the main concentration results, i.e. Theorems 3.2, 3.5, 3.8, respectively under Theorems C.10, C.11, C.12.

Theorem C.10 (Concentration inequality for RFSF kernel). Let $\mathfrak{X} \subset \mathbb{R}^d$ be a compact and convex set with diameter $|\mathfrak{X}|$, and $\mathfrak{X}_{\Delta} := \{\mathbf{x} - \mathbf{y} : \mathbf{x}, \mathbf{y} \in \mathfrak{X}\}$. Let $\mathbf{k} : \mathbb{R}^d \to \mathbb{R}$ be a continuous, bounded, translation-invariant kernel with spectral measure Λ , which satisfies for some S, R > 0 that

$$\mathbb{E}_{\mathbf{w}\sim\Lambda}\left[|w_i|^{2k}\right] \leq \frac{k!S^2R^{k-2}}{2} \quad for all \ i \in [d] \ and \ k \geq 2.$$

$$\tag{45}$$

Then, the following quantities are finite: $\sigma_{\Lambda}^2 := \mathbb{E}_{\mathbf{w} \sim \Lambda} [\|\mathbf{w}\|_2^2], L := \|\mathbb{E}_{\mathbf{w} \sim \Lambda} [\|\mathbf{w}\|_2^T] \|_2^{1/2}, E_{i,j} := \mathbb{E}_{\mathbf{w} \sim \Lambda} [\|\mathbf{w}\|_2^T] \|\mathbf{w}\|_2^T$ and $D_{i,j} := \sup_{\mathbf{z} \in \mathcal{X}_{\Delta}} \|\nabla \left[\frac{\partial^2 \mathbf{k}(\mathbf{z})}{\partial z_i \partial z_j}\right] \|_2$ for $i, j \in [d]$. Further, for any maximal sequence 1-variation V > 0, and signature level $m \in \mathbb{Z}_+$, it holds for the level- $m \operatorname{RFSF}$ kernel $\tilde{k}_{\operatorname{Sig}_m} : \mathcal{X}_{\operatorname{seq}} \times \mathcal{X}_{\operatorname{seq}} \to \mathbb{R}$ defined as in (14) and the signature kernel $k_{\operatorname{Sig}_m} : \mathcal{X}_{\operatorname{seq}} \times \mathcal{X}_{\operatorname{seq}} \to \mathbb{R}$ defined as in (12) for $\epsilon > 0$ that

$$\mathbb{P}\left[\sup_{\substack{\mathbf{x},\mathbf{y}\in\mathcal{X}_{\mathrm{seq}}\\||\mathbf{x}||_{1-\mathrm{var}},||\mathbf{y}||_{1-\mathrm{var}}\leq V}}\left|\mathsf{k}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y})-\tilde{\mathsf{k}}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y})\right|\geq\epsilon\right]\leq$$

$$\leq m \begin{cases} \left(C_{d,\chi} \left(\frac{\beta_{d,m,V}}{\epsilon} \right)^{\frac{d}{d+1}} + d \right) \exp\left(-\frac{\tilde{d}}{2(d+1)(S^2+R)} \left(\frac{\epsilon}{\beta_{d,m,V}} \right)^2 \right) & \text{for } \epsilon < \beta_{d,m,V} \\ \left(C_{d,\chi} \left(\frac{\beta_{d,m,V}}{\epsilon} \right)^{\frac{d}{(d+1)m}} + d \right) \exp\left(-\frac{\tilde{d}}{2(d+1)(S^2+R)} \left(\frac{\epsilon}{\beta_{d,m,V}} \right)^{\frac{1}{m}} \right) & \text{for } \epsilon \geq \beta_{d,m,V}, \end{cases}$$

where $C_{d,\chi} := 2^{\frac{1}{d+1}} 16 |\chi|^{\frac{d}{d+1}} \sum_{i,j=1}^{d} (D_{i,j} + E_{i,j})^{\frac{d}{d+1}} and \beta_{d,m,V} := m (2V^2 (L^2 \vee 1) (\sigma_{\Lambda}^2 \vee d))^m.$

Proof. Finite quantities. To start off with, due to (45) with m = 2 and Jensen's inequality (Lemma A.1), we get

$$\mathbb{E}\big[|w_i|^2\big] \le \mathbb{E}^{1/2}\big[|w_i|^4\big] < \infty \quad \text{for all } i \in [d].$$

Hence, by linearity of the expectation $\sigma_{\Lambda}^2 = \mathbb{E}[\|\mathbf{w}\|^2] = \sum_{i=1}^d \mathbb{E}[w_i^2] < \infty$. Next, due to Hölder's inequality (Lemma A.2), it holds that

$$\mathbb{E}[w_i w_j] \le \mathbb{E}[|w_i w_j|] \le \mathbb{E}^{1/2}[|w_i|^2] \mathbb{E}^{1/2}[|w_j|^2] < \infty \quad \text{for all } i, j \in [d],$$

therefore $L < \infty$. Further, applying Hölder's inequality twice followed by Jensen's inequality,

$$\mathbb{E}[|w_{i}w_{j}w_{k}|] \leq \mathbb{E}^{2/3}[|w_{i}w_{j}|^{3/2}]\mathbb{E}^{1/3}[|w_{k}|^{3}] \leq \mathbb{E}^{1/3}[|w_{i}|^{3}]\mathbb{E}^{1/3}[|w_{j}|^{3}]\mathbb{E}^{1/3}[|w_{k}|^{3}] \\ \leq \mathbb{E}^{1/6}[|w_{i}|^{6}]\mathbb{E}^{1/6}[|w_{j}|^{6}]\mathbb{E}^{1/6}[|w_{k}|^{6}] < \infty \quad \text{for all } i, j, k \in [d],$$

$$(46)$$

which is finite due to (45) with m = 3. Now, because of the $\ell^1 - \ell^2$ norm inequality,

$$\mathbb{E}\left[\left|w_{i}w_{j}\right|\|\mathbf{w}\|_{2}\right] \leq \mathbb{E}\left[\left|w_{i}w_{j}\right|\|\mathbf{w}\|_{1}\right] = \sum_{k=1}^{d} \mathbb{E}\left[\left|w_{i}w_{j}w_{k}\right|\right] < \infty,$$

hence $\bar{E} < \infty$. Next, as per [70, Thm. 1.2.1.(iii)], as (46) holds for all $i, j, k \in [d]$, k is 3-times continuously differentiable, which combined with the compactness of \mathcal{X} , hence that of \mathcal{X}_{Δ} , gives $\sup_{\mathbf{z} \in \mathcal{X}_{\Delta}} \left| \frac{\partial^3 \mathbf{k}(\mathbf{z})}{\partial z_i \partial z_j \partial z_k} \right| < \infty$. Finally, from the $\ell^1 - \ell^2$ inequality again, we get that

$$\sup_{\mathbf{z}\in\mathcal{X}_{\Delta}}\left\|\nabla\left[\frac{\partial^{2}\mathsf{k}(\mathbf{z})}{\partial z_{i}\partial z_{j}}\right]\right\|_{2} \leq \sup_{\mathbf{z}\in\mathcal{X}_{\Delta}}\left\|\nabla\left[\frac{\partial^{2}\mathsf{k}(\mathbf{z})}{\partial z_{i}\partial z_{j}}\right]\right\|_{1} \leq \sum_{k=1}^{d} \sup_{\mathbf{z}\in\mathcal{X}_{\Delta}}\left|\frac{\partial^{3}\mathsf{k}(\mathbf{z})}{\partial z_{i}\partial z_{j}\partial z_{k}}\right| < \infty,$$

which shows the finiteness of \overline{D} . This finishes showing that the stated quantities are finite.

Splitting the bound. To start proving inequality our main inequality, first note that as per Example C.2, k is *L*-Lipschitz (see Def. C.1). Hence, Lemma C.9 yields that

$$\sup_{\substack{\mathbf{x},\mathbf{y}\in\mathcal{X}_{\mathrm{seq}}\\\|\mathbf{x}\|_{1-\mathrm{var}},\|\mathbf{y}\|_{1-\mathrm{var}}\leq V}} \left| \tilde{\mathbf{k}}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y}) - \tilde{\mathbf{k}}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y}) \right| \\
\leq V^{2m} \sum_{k=1}^{m} \frac{L^{2(m-k)}}{\tilde{d}^{k-1}((k-1)!)^{2}} \sup_{\mathbf{s},\mathbf{t}\in\mathcal{X}} \left\| \partial_{\mathbf{s},\mathbf{t}}^{2} \tilde{\mathbf{k}}_{k}(\mathbf{s},\mathbf{t}) - \partial_{\mathbf{s},\mathbf{t}}^{2} \mathbf{k}(\mathbf{s},\mathbf{t}) \right\|_{2} \prod_{p=1}^{k-1} \left\| \mathbf{W}^{(p)} \right\|_{2}^{2}. \quad (47)$$

We bound the summand in the previous line in probability for each $k \in [m]$. For brevity, denote $\alpha_{m,k} \coloneqq \frac{V^{2m}L^{2(m-k)}}{((k-1)!)^2}$, and first consider the case $k \ge 2$, so that we have

$$P_{k}(\epsilon) \coloneqq \mathbb{P}\left[\frac{\alpha_{m,k}}{\tilde{d}^{k-1}} \left\|\mathbf{W}^{(1)}\right\|_{2}^{2} \cdots \left\|\mathbf{W}^{(k-1)}\right\|_{2}^{2} \sup_{\mathbf{s},\mathbf{t}\in\mathcal{X}} \left\|\partial_{\mathbf{s},\mathbf{t}}\tilde{\mathsf{k}}_{k}(\mathbf{s},\mathbf{t}) - \partial_{\mathbf{s},\mathbf{t}}\mathsf{k}(\mathbf{s},\mathbf{t})\right\|_{2} \ge \epsilon\right]$$

$$\overset{(\mathbf{a})}{\leq} \mathbb{P}\left[\alpha_{m,k} \left(\frac{\left\|\mathbf{W}^{(1)}\right\|_{2}^{2} + \ldots + \left\|\mathbf{W}^{(k-1)}\right\|_{2}^{2}}{\tilde{d}(k-1)}\right)^{k-1} \sup_{\mathbf{s}, \mathbf{t} \in \mathcal{X}} \left\|\partial_{\mathbf{s}, \mathbf{t}} \tilde{\mathbf{k}}_{k}(\mathbf{s}, \mathbf{t}) - \partial_{\mathbf{s}, \mathbf{t}} \mathbf{k}(\mathbf{s}, \mathbf{t})\right\|_{2} \geq \epsilon\right]$$

$$\overset{(\mathbf{b})}{=} \mathbb{P}\left[\underbrace{\underbrace{\left\|\mathbf{W}^{(1)}\right\|_{2}^{2} + \ldots + \left\|\mathbf{W}^{(k-1)}\right\|_{2}^{2}}_{(A_{k})} \sup_{\mathbf{s}, \mathbf{t} \in \mathcal{X}} \left\|\partial_{\mathbf{s}, \mathbf{t}} \tilde{\mathbf{k}}_{k}(\mathbf{s}, \mathbf{t}) - \partial_{\mathbf{s}, \mathbf{t}} \mathbf{k}(\mathbf{s}, \mathbf{t})\right\|_{2}^{\frac{1}{k-1}}}_{(B_{k})} \geq \left(\frac{\epsilon}{\alpha_{m, k}}\right)^{\frac{1}{k-1}}\right],$$

where in (a) we used the arithmetic-geometric mean inequality, and in (b) we divided both sides by $\alpha_{m,k}$ and took the (k-1)th root. Further, setting $t \coloneqq \left(\frac{\epsilon}{\alpha_{m,k}}\right)^{\frac{1}{k-1}}$, we have for $\gamma > 0$

$$P_{k}(\epsilon) \leq \qquad \mathbb{P}\left[A_{k} \cdot B_{k} \geq t\right] \stackrel{(c)}{\leq} \mathbb{P}\left[\left(A_{k} - \gamma\right)B_{k} \geq \frac{t}{2}\right] + \mathbb{P}\left[B_{k} \geq \frac{t}{2\gamma}\right] \\ \stackrel{(d)}{\leq} \qquad \inf_{\tau > 0}\left\{\mathbb{P}\left[A_{k} - \gamma \geq \frac{\tau}{2}\right] + \mathbb{P}\left[B_{k} \geq \frac{t}{\tau}\right]\right\} + \mathbb{P}\left[B_{k} \geq \frac{t}{2\gamma}\right], \qquad (48)$$

where in (c) we added and subtracted γB_k and applied a union bound, while in (d) we combined a union bound with the relation $\{X Y \ge \epsilon\} \subseteq \{X \ge \tau\} \bigcup \{Y \ge \epsilon/\tau\}$ which holds for any $\tau > 0$.

Our aim is now to obtain good probabilistic bounds on A_k and B_k to use in (48) with the specific choice of $\gamma = \mathbb{E}[A_k] = \sigma_{\Lambda}^2$.

Bounding A_k . By the inequality between the spectral and Frobenius norms, we have

$$\sum_{p=1}^{k-1} \left\| \mathbf{W}^{(p)} \right\|_{2}^{2} \leq \sum_{p=1}^{k-1} \left\| \mathbf{W}^{(p)} \right\|_{F}^{2} = \sum_{i=1}^{d} \sum_{p=1}^{k-1} \sum_{j=1}^{\tilde{d}} \left(w_{i,j}^{(p)} \right)^{2}.$$

Now note that $w_{i,j}^{(p)}$ are i.i.d. copies of the i^{th} marginal of Λ for all $j \in [\tilde{d}]$ and $p \in [k-1]$, so that via (45) $(w_{i,j}^{(p)})^2$ satisfies the Bernstein moment condition

$$\mathbb{E}\left[\left(w_{i,j}^{(p)}\right)^{2k}\right] \leq \frac{k!S^2R^{k-2}}{2} \quad \text{for all } i \in [d], j \in [\tilde{d}], p \in [m], k \geq 2.$$

Hence, we may apply the Bernstein inequality from Theorem A.5 so that for $i \in [d]$

$$\mathbb{P}\left[\frac{1}{\tilde{d}(k-1)}\sum_{p=1}^{k-1}\sum_{j=1}^{\tilde{d}}\left(w_{i,j}^{(p)}\right)^2 - \sigma_i^2 \ge \epsilon\right] \le \exp\left(\frac{-\tilde{d}(k-1)\epsilon^2}{2(S^2 + R\epsilon)}\right),$$

where $\sigma_i^2 = \mathbb{E}_{\mathbf{w} \sim \Lambda}[w_i^2]$. Combining these bounds for all $i \in [d]$ and denoting $\sigma_{\Lambda}^2 = \sum_{i=1}^d \sigma_i^2$,

$$\mathbb{P}\left[A_{k}-\sigma_{\Lambda}^{2}\geq\epsilon\right]=\mathbb{P}\left[\frac{1}{\tilde{d}(k-1)}\sum_{i=1}^{d}\sum_{p=1}^{k-1}\sum_{j=1}^{\tilde{d}}\left(w_{i,j}^{(p)}\right)^{2}-\sigma_{\Lambda}^{2}\geq\epsilon\right]$$

$$\leq\sum_{i=1}^{d}\mathbb{P}\left[\frac{1}{\tilde{d}(k-1)}\sum_{p=1}^{k-1}\sum_{j=1}^{\tilde{d}}\left(w_{i,j}^{(p)}\right)^{2}-\sigma_{i}^{2}\geq\frac{\epsilon}{d}\right]$$

$$\leq d\exp\left(\frac{-\tilde{d}(k-1)\left(\frac{\epsilon}{d}\right)^{2}}{2\left(S^{2}+R\frac{\epsilon}{d}\right)}\right).$$
(49)

Hence, we have the required probabilistic bound for the term in (48) containing A_k . Bounding B_k . One can bound the spectral norm by the max norm so that

$$B_{k}^{k-1} = \sup_{\mathbf{s},\mathbf{t}\in\mathcal{X}} \left\| \partial_{\mathbf{s},\mathbf{t}}\tilde{\kappa}_{k}(\mathbf{s},\mathbf{t}) - \partial_{\mathbf{s},\mathbf{t}}\boldsymbol{\kappa}(\mathbf{s},\mathbf{t}) \right\|_{2} \leq \sup_{\mathbf{s},\mathbf{t}\in\mathcal{X}} \left\| \partial_{\mathbf{s},\mathbf{t}}\tilde{\kappa}_{k}(\mathbf{s},\mathbf{t}) - \partial_{\mathbf{s},\mathbf{t}}\boldsymbol{\kappa}(\mathbf{s},\mathbf{t}) \right\|_{\max} \\ = \max_{i,j=1,\dots,d} \sup_{\mathbf{s},\mathbf{t}\in\mathcal{X}} \left| \frac{\partial^{2}\tilde{\kappa}_{k}(\mathbf{s},\mathbf{t})}{\partial s_{i}\partial t_{j}} - \frac{\partial^{2}\boldsymbol{\kappa}(\mathbf{s},\mathbf{t})}{\partial s_{i}\partial t_{j}} \right|.$$
(50)

Let $i, j \in [d]$ and denote $E_{i,j} \coloneqq \mathbb{E}_{\mathbf{w} \sim \Lambda} [|w_i w_j| \|\mathbf{w}\|_2]$ and $D_{i,j} \coloneqq \sup_{\mathbf{z} \in \mathcal{X}_{\Delta}} \|\nabla [\partial^{\mathbf{e}_i, \mathbf{e}_j} \mathsf{k}(\mathbf{z})]\|_2$, which are finite as previously shown. Due to Hölder's inequality (Lemma A.2) and (45),

$$\mathbb{E}\left[\left|w_{i}w_{j}\right|^{m}\right] \leq \mathbb{E}^{1/2}\left[w_{i}^{2m}\right]\mathbb{E}^{1/2}\left[w_{j}^{2m}\right] < \infty.$$

Recall that k is 3-times continuously differentiable, so that the conditions required by Theorem B.1 are satisfied, that we now call to our aid in controlling the RFF kernel derivatives,

$$\mathbb{P}\left[\sup_{\mathbf{s},\mathbf{t}\in\mathcal{X}}\left|\frac{\partial^{2}\tilde{k}_{k}(\mathbf{s},\mathbf{t})}{\partial s_{i}\partial t_{j}}-\frac{\partial^{2}k(\mathbf{s},\mathbf{t})}{\partial s_{i}\partial t_{j}}\right|\geq\epsilon\right]\leq16C_{d,\mathcal{X},i,j}^{\prime}\epsilon^{-\frac{d}{d+1}}\exp\left(\frac{-\tilde{d}\epsilon^{2}}{4(d+1)(2S^{2}+R\epsilon)}\right),$$

where we defined $C'_{d,\mathfrak{X},i,j} \coloneqq (|\mathfrak{X}|(D_{i,j} + E_{i,j}))^{\frac{d}{d+1}}$. Hence, noting that the max satisfies the relation $\{\max_i \xi_i \ge \epsilon\} = \bigcup_i \{\xi_i \ge \epsilon\}$ and union bounding (50) in probability, we get that

$$\mathbb{P}\left[B_{k}^{k-1} \geq \epsilon\right] \leq \sum_{i,j=1}^{d} \mathbb{P}\left[\sup_{\mathbf{s},\mathbf{t}\in\mathcal{X}} \left|\frac{\partial^{2}\tilde{\mathsf{k}}_{k}(\mathbf{s},\mathbf{t})}{\partial s_{i}\partial t_{j}} - \frac{\partial^{2}\mathsf{k}(\mathbf{s},\mathbf{t})}{\partial s_{i}\partial t_{j}}\right| \geq \epsilon\right]$$
$$\leq 16C_{d,\mathcal{X}}^{\prime} \epsilon^{-\frac{d}{d+1}} \exp\left(\frac{-\tilde{d}\epsilon^{2}}{4(d+1)(2S^{2}+R\epsilon)}\right),\tag{51}$$

where we denote $C'_{d,\mathfrak{X}} \coloneqq \sum_{i,j=1}^{d} C'_{d,\mathfrak{X},i,j} = |\mathfrak{X}|^{\frac{d}{d+1}} \sum_{i,j=1}^{d} (D_{i,j} + E_{i,j})^{\frac{d}{d+1}}$. *Putting it together.* Now that we have our bounds for A_k and B_k , we put everything

together, that is, plug the bounds (49) and (51) into (48), so that we get

$$\begin{split} P_{k}(\epsilon) &\leq \inf_{\tau>0} \left\{ d \exp\left(\frac{-\tilde{d}(k-1)\left(\frac{\tau}{2d}\right)^{2}}{2\left(S^{2}+R\frac{\tau}{2d}\right)}\right) + 16C_{d,\mathcal{X}}'\left(\frac{\tau}{t}\right)^{\frac{d(k-1)}{d+1}} \exp\left(\frac{-\tilde{d}\left(\frac{t}{\tau}\right)^{2(k-1)}}{4(d+1)\left(2S^{2}+R\left(\frac{t}{\tau}\right)^{k-1}\right)}\right) \right\} \\ &+ 16C_{d,\mathcal{X}}'\left(\frac{2\sigma_{\Lambda}^{2}}{t}\right)^{\frac{d(k-1)}{d+1}} \exp\left(\frac{-\tilde{d}\left(\frac{t}{2\sigma_{\Lambda}^{2}}\right)^{2(k-1)}}{4(d+1)\left(2S^{2}+R\left(\frac{t}{2\sigma_{\Lambda}^{2}}\right)^{k-1}\right)}\right) \right\} \\ &\stackrel{(e)}{\leq} d \exp\left(\frac{-\tilde{d}(k-1)\left(\frac{t^{\frac{k-1}{2}}}{2d}\right)^{2}}{2\left(S^{2}+R\frac{t^{\frac{k-1}{2}}}{2d}\right)}\right) + 16C_{d,\mathcal{X}}'\left(\frac{1}{t}\right)^{\frac{d(k-1)}{d+1k}} \exp\left(\frac{-\tilde{d}t\frac{2(k-1)}{k}}{4(d+1)\left(2S^{2}+Rt\frac{k-1}{k}\right)}\right) \right) \\ &+ 16C_{d,\mathcal{X}}'\left(\frac{2\sigma_{\Lambda}^{2}}{t}\right)^{\frac{d(k-1)}{d+1}} \exp\left(\frac{-\tilde{d}\left(\frac{t}{2\sigma_{\Lambda}^{2}}\right)^{2(k-1)}}{4(d+1)\left(2S^{2}+R\left(\frac{t}{2\sigma_{\Lambda}^{2}}\right)^{k-1}\right)}\right) \right) \\ &\stackrel{(f)}{=} d \exp\left(\frac{-\tilde{d}(k-1)\left(\frac{(\epsilon/a_{m,k})^{\frac{1}{k}}}{2d}\right)^{2}}{2\left(S^{2}+R\frac{(\epsilon/a_{m,k})^{\frac{1}{k}}}{2d}\right)} + 16C_{d,\mathcal{X}}'\left(\frac{a_{m,k}}{\epsilon}\right)^{\frac{d}{(d+1)k}} \exp\left(\frac{-\tilde{d}\left(\frac{e}{a_{m,k}}\right)^{\frac{2}{k}}}{4(d+1)\left(2S^{2}+R\left(\frac{e}{a_{m,k}}\right)^{\frac{1}{k}}}\right)\right) \end{split}$$

$$+16C_{d,\mathfrak{X}}^{\prime}\left(\frac{\alpha_{m,k}(2\sigma_{\Lambda}^{2})^{k-1}}{\epsilon}\right)^{\frac{d}{d+1}}\exp\left(\frac{-\tilde{d}\left(\frac{\epsilon}{\alpha_{m,k}(2\sigma_{\Lambda}^{2})^{k-1}}\right)^{2}}{4(d+1)\left(2S^{2}+R\frac{\epsilon}{\alpha_{m,k}(2\sigma_{\Lambda}^{2})^{k-1}}\right)}\right),$$

where (a) follows from substituting (51) and (49) into (48) with the choice of $\gamma = \sigma_{\Lambda}^2$, (b) from choosing $\tau = t^{\frac{k-1}{k}}$, and (c) from putting back $t = (\epsilon/\alpha_{m,k})^{\frac{1}{k-1}}$. Note that the previous applies for all $k \ge 2$. For k = 1, we have by (51) that

$$P_{1}(\epsilon) = \mathbb{P}\left[\sup_{\mathbf{s},\mathbf{t}\in\mathcal{X}} \left\|\partial_{\mathbf{s},\mathbf{t}}\tilde{\mathbf{k}}_{1}(\mathbf{s},\mathbf{t}) - \partial_{\mathbf{s},\mathbf{t}}\mathbf{k}(\mathbf{s},\mathbf{t})\right\|_{2} \ge \frac{\epsilon}{\alpha_{m,1}}\right]$$
$$\leq 16C_{d,\mathcal{X}}'\left(\frac{\alpha_{m,1}}{\epsilon}\right)^{\frac{d}{d+1}}\exp\left(\frac{-\tilde{d}\left(\frac{\epsilon}{\alpha_{m,1}}\right)^{2}}{4(d+1)\left(2S^{2}+R\frac{\epsilon}{\alpha_{m,1}}\right)}\right).$$

Combining and simplifying. We can now combine the bounds for P_1, \ldots, P_m into (47),

$$\begin{split} \mathbb{P}\left[\sup_{\substack{\mathbf{x},\mathbf{y}\in\mathcal{X}_{seg}\\||\mathbf{x}||_{1\leq urr}||\mathbf{y}||_{1\leq urr}\leq V}} |\tilde{k}_{Sig_{m}}(\mathbf{x},\mathbf{y})-k_{Sig_{m}}(\mathbf{x},\mathbf{y})| \geq \epsilon\right] \leq \sum_{k=1}^{m} P_{k}\left(\frac{\epsilon}{m}\right) \\ & \left(\frac{e}{2}2^{\frac{1}{d+1}}8C'_{d,\mathcal{X}}\sum_{k=1}^{m}\left(\frac{2^{k}\sigma_{\Lambda}^{2\,k-1}m\alpha_{m,k}}{\epsilon}\right)^{\frac{d}{d+1}}\exp\left(-\frac{\tilde{d}}{2(d+1)}\cdot\frac{\left(\frac{2^{k}\sigma_{\Lambda}^{2(k-1)}m\alpha_{m,k}}{S^{2}+R\frac{2^{k}\sigma_{\Lambda}^{2(k-1)}m\alpha_{m,k}}\right)^{2}}\right) \\ & +2^{\frac{1}{d+1}}8C'_{d,\mathcal{X}}\sum_{k=2}^{m}\left(\frac{2^{k}m\alpha_{m,k}}{\epsilon}\right)^{\frac{d}{d+1}k}\exp\left(-\frac{\tilde{d}}{2(d+1)}\cdot\frac{\left(\frac{e}{2^{k}\sigma_{\Lambda}^{2(k-1)}m\alpha_{m,k}}\right)^{\frac{1}{k}}}{S^{2}+R\left(\frac{e}{2^{k}\sigma_{\Lambda}^{2(k-1)}m\alpha_{m,k}}\right)^{\frac{1}{k}}}\right) \\ & +d\sum_{k=2}^{M}\exp\left(-\frac{\tilde{d}(k-1)}{2}\frac{\left(\frac{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}{\epsilon}\right)^{\frac{d}{d+1}}\exp\left(-\frac{\tilde{d}}{2(d+1)}\cdot\frac{\left(\frac{e}{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}\right)^{2}}{S^{2}+R\left(\frac{e}{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}\right)^{\frac{d}{d+1}}}\right) \\ & +2^{\frac{1}{d+1}}8C'_{d,\mathcal{X}}\sum_{k=1}^{m}\left(\frac{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}{\epsilon}\right)^{\frac{1}{d+1}}\exp\left(-\frac{\tilde{d}}{2(d+1)}\cdot\frac{\left(\frac{e}{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}\right)^{\frac{1}{k}}\right) \\ & +2^{\frac{1}{d+1}}8C'_{d,\mathcal{X}}\sum_{k=1}^{m}\left(\frac{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}{\epsilon}\right)^{\frac{1}{d+1}}\exp\left(-\frac{\tilde{d}}{2(d+1)}\cdot\frac{\left(\frac{e}{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}\right)^{\frac{1}{k}}\right) \\ & +d\sum_{k=1}^{m}\exp\left(-\frac{\tilde{d}}{2(d+1)}\frac{\left(\frac{e}{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}\right)^{\frac{1}{d+1}}\exp\left(-\frac{\tilde{d}}{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}\right)^{\frac{1}{k}}\right) \\ & (1)2^{\frac{1}{d+1}}8C'_{d,\mathcal{X}}\sum_{k=1}^{m}\left(\frac{\beta_{d,m,V}}{2(d+1)}\frac{1}{S^{2}+R\left(\frac{e}{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}\right)^{\frac{1}{d+1}}}\exp\left(-\frac{\tilde{d}}{2(d+1)}\frac{e}{S^{2}+R\frac{e}{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}}\right)^{\frac{1}{k}}\right) \\ & (1)2^{\frac{1}{d+1}}8C'_{d,\mathcal{X}}\sum_{k=1}^{m}\left(\frac{\beta_{d,m,V}}{2(d+1)}\frac{1}{S^{2}+R\left(\frac{e}{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}\right)^{\frac{1}{d+1}}}}{S^{2}+R\left(\frac{e}{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}\right)^{\frac{1}{k}}}\right) \\ & (1)2^{\frac{1}{d+1}}8C'_{d,\mathcal{X}}\sum_{k=1}^{m}\left(\frac{\beta_{d,m,V}}{2(d+1)}\frac{1}{S^{2}+R\left(\frac{e}{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}\right)^{\frac{1}{d+1}}}}}{S^{2}+R\left(\frac{e}{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}\right)^{\frac{1}{k}}}}\right) \\ & (1)2^{\frac{1}{d+1}}8C'_{d,\mathcal{X}}\sum_{k=1}^{m}\left(\frac{\beta_{d,m,V}}{2(d+1)}\frac{1}{S^{2}+R\left(\frac{e}{2^{k}(\sigma_{\Lambda}^{2}\vee d)^{k}m\alpha_{m,k}}\right)^{\frac{1}{k}}}}}{S^{2}+R\left(\frac{e}{2^{k}(\sigma_{\Lambda$$

$$+2^{\frac{1}{d+1}}8C'_{d,\mathcal{X}}\sum_{k=1}^{m}\left(\frac{\beta_{d,m,V}}{\epsilon}\right)^{\frac{d}{(d+1)k}}\exp\left(-\frac{\tilde{d}}{2(d+1)}\frac{\left(\frac{\epsilon}{\beta_{d,m,V}}\right)^{\frac{2}{k}}}{S^{2}+R\left(\frac{\epsilon}{\beta_{d,m,V}}\right)^{\frac{1}{k}}}\right)$$
$$+d\sum_{k=1}^{m}\exp\left(-\frac{\tilde{d}}{2(d+1)}\frac{\left(\frac{\epsilon}{\beta_{d,m,V}}\right)^{\frac{2}{k}}}{S^{2}+R\left(\frac{\epsilon}{\beta_{d,m,V}}\right)^{\frac{1}{k}}}\right),$$

where (g) follows from rearranging the expressions from (f), while (h) from unifying the coefficients and that $1 \le d \le \max(\sigma_{\Lambda}^2, d)$ and $f(x) = x^2/(a+bx)$ is monotonically increasing in x on the positive half-line for a, b > 0, while (f) from $\alpha_{m,k} = V^{2m}L^{2(m-k)}/((k-1)!)^2 \le (VL)^{2m}$, using that f(x) is increasing, and defining $\beta_{d,m,V} \coloneqq m \left(2V^2(L^2 \lor 1)(\sigma_{\Lambda}^2 \lor d)\right)^m$.

 $(VL)^{2m}$, using that f(x) is increasing, and defining $\beta_{d,m,V} \coloneqq m \left(2V^2(L^2 \vee 1)(\sigma_{\Lambda}^2 \vee d)\right)^m$. *Conclusion.* Finally, we split the bound into two cases: the first case is if the error is big, i.e. $\epsilon \ge \beta_{d,m,V} = m \left(2V^2(L^2 \vee 1)(\sigma_{\Lambda}^2 \vee d)\right)^m$, when we decrease all the exponents to $\frac{1}{m}$,

$$\mathbb{P}\left[\sup_{\substack{\mathbf{x},\mathbf{y}\in\mathcal{X}_{\mathrm{seq}}\\\|\mathbf{x}\|_{1-\mathrm{var}},\|\mathbf{y}\|_{1-\mathrm{var}}\leq V}}\left|\tilde{\mathbf{k}}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y})-\mathbf{k}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y})\right|\geq\epsilon\right]$$
$$\leq m\cdot\left(2^{\frac{1}{d+1}}16C'_{d,\mathcal{X}}\left(\frac{\beta_{d,m,V}}{\epsilon}\right)^{\frac{d}{(d+1)m}}+d\right)\exp\left(-\frac{\tilde{d}}{2(d+1)}\frac{\left(\frac{\epsilon}{\beta_{d,m,V}}\right)^{\frac{2}{m}}}{S^{2}+R\left(\frac{\epsilon}{\beta_{d,m,V}}\right)^{\frac{1}{m}}}\right)$$

The other when the error is small, i.e. $\epsilon < \beta_{d,m,V}$, when we increase all the exponents to 1

$$\mathbb{P}\left[\sup_{\substack{\mathbf{x},\mathbf{y}\in\mathcal{X}_{\mathrm{seq}}\\||\mathbf{x}||_{1-\mathrm{var}},||\mathbf{y}||_{1-\mathrm{var}}\leq V}}\left|\tilde{k}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y})-k_{\mathrm{Sig}}(\mathbf{x},\mathbf{y})\right|\geq\epsilon\right]$$

$$\leq m\cdot\left(2^{\frac{1}{d+1}}16C'_{d,\mathcal{X}}\left(\frac{\beta_{d,m,V}}{\epsilon}\right)^{\frac{d}{d+1}}+d\right)\exp\left(-\frac{\tilde{d}}{2(d+1)}\frac{\left(\frac{\epsilon}{\beta_{d,m,V}}\right)^{2}}{S^{2}+R\left(\frac{\epsilon}{\beta_{d,m,V}}\right)}\right).$$

The claimed estimate follows by denoting $C_{d,\mathcal{X}} \coloneqq 2^{\frac{1}{d+1}} 16C'_{d,\mathcal{X}}$ and simplifying.

Next, we prove Theorem 3.5 to show an approximation bound for the RFSF-DP kernel.

Theorem C.11 (Concentration inequality for RFSF-DP kernel). Let $k : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ be a continuous, bounded, translation-invariant kernel with spectral measure Λ , which satisfies for some S, R > 0 that

$$\mathbb{E}_{\mathbf{w}\sim\Lambda}\left[|w_i|^{2k}\right] \leq \frac{k!S^2R^{k-2}}{2} \quad for all \ i \in [d] \ and \ k \geq 2.$$
(52)

Then, for signature level $m \in \mathbb{Z}_+$ *and* $\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq}$ *, it holds for* $\epsilon > 0$ *that:*

$$\mathbb{P}\left[\left|\tilde{\mathsf{k}}_{\operatorname{Sig}_{m}}^{\mathsf{DP}}(\mathbf{x},\mathbf{y})-\mathsf{k}_{\operatorname{Sig}_{m}}(\mathbf{x},\mathbf{y})\right| \geq \epsilon\right] \leq 2 \exp\left(-\frac{1}{4} \min\left\{ \begin{array}{c} \left(\frac{\sqrt{d}\epsilon}{2C_{d,m}\|\mathbf{x}\|_{1-\operatorname{var}}^{m}}\|\mathbf{y}\|_{1-\operatorname{var}}^{m}\right)^{2} \\ \left(\frac{d}{\sqrt{8}C_{d,m}\|\mathbf{x}\|_{1-\operatorname{var}}^{m}}\|\mathbf{y}\|_{1-\operatorname{var}}^{m}\right)^{\frac{1}{m}} \end{array}\right),$$

where the absolute constant $C_{d,m} > 0$ satisfies that

$$C_{d,m} \le \sqrt{8}e^4(2\pi)^{1/4}e^{1/24}(4e^3/m)^m ((2d\max(S,R))^m + (L^2/\ln 2)^m).$$

Proof. Let $\hat{k}_{\text{Sig}_m}^{(1)}, \dots, \hat{k}_{\text{Sig}_m}^{(\tilde{d})}$ be independent copies of the RFSF kernel, each with internal RFF sample size $\hat{d} = 1$, such that $\tilde{k}_{\text{Sig}_m}^{\text{DP}} = \frac{1}{\tilde{d}} \sum_{k=1}^{\tilde{d}} \hat{k}_{\text{Sig}_m}^{(k)}$. Our goal is to call Theorem A.13 with $\alpha = \frac{1}{m}$, and therefore, we compute an upper bound on the $\Psi_{1/m}$ -norm of $\tilde{k}_{\text{Sig}_m}^{(k)}(\mathbf{x}, \mathbf{y}) - k_{\text{Sig}_m}(\mathbf{x}, \mathbf{y})$ for all $k \in [\tilde{d}]$; for a definition of the α -exponential norm, see Definition A.7.

By Lemma C.6, it holds for any $\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq}$ and $k \in [m]$ that

$$\left| \hat{\mathsf{k}}_{\text{Sig}_{m}}^{(k)}(\mathbf{x},\mathbf{y}) \right| \leq \frac{\left(||\mathbf{x}||_{1-\text{var}} \, \left\| \mathbf{y} \right\|_{1-\text{var}} \right)^{m}}{(m!)^{2}} \left\| \mathbf{w}_{k}^{(1)} \right\|_{2}^{2} \cdots \left\| \mathbf{w}_{k}^{(m)} \right\|_{2}^{2},$$

where $\mathbf{w}_k^{(1)}, \ldots, \mathbf{w}_k^{(m) \text{ i.i.d.}} \wedge \text{ are the random weights that parametrize } \hat{k}_{\text{Sig}_m}^{(k)}$ for all $k \in [\tilde{d}]$. Now, calling Lemma A.11 with $\alpha = 1$ yields that, due to (52), the following holds:

$$\left\| w_{k,i}^{(p)^2} \right\|_{\Psi_1} \le 2 \max(S, R) \text{ for all } i \in [d], k \in [\tilde{d}], p \in [m].$$

Note that for $\alpha = 1$, $\|\cdot\|_{\Psi_{\alpha}}$ satisfies the triangle inequality (see Lemma A.9), and hence

$$\left\| \left\| \mathbf{w}_{k}^{(p)} \right\|_{2}^{2} \right\|_{\Psi_{1}} \leq \sum_{i=1}^{d} \left\| w_{k,i}^{(p)^{2}} \right\|_{\Psi_{1}} \leq 2d \max(S,R) \quad \text{for all } k \in [\tilde{d}], p \in [m].$$

As $\|\cdot\|_{\Psi_{\alpha}}$ is positive homogenous and satisfies a Hölder-type inequality (see Lemma A.10):

$$\begin{aligned} \left\| \hat{\mathsf{k}}_{\text{Sig}_{m}}^{(k)}(\mathbf{x}, \mathbf{y}) \right\|_{\Psi_{1/m}} & \stackrel{\text{(a)}}{\leq} \frac{\left(\|\mathbf{x}\|_{1-\text{var}} \|\mathbf{y}\|_{1-\text{var}} \right)^{m}}{(m!)^{2}} \left\| \left\| \mathbf{w}_{k}^{(1)} \right\|_{2}^{2} \right\|_{\Psi_{1}} \cdots \left\| \left\| \mathbf{w}_{k}^{(m)} \right\|_{2}^{2} \right\|_{\Psi_{1}} \\ & \stackrel{\text{(b)}}{\leq} \frac{\left(2d \|\mathbf{x}\|_{1-\text{var}} \|\mathbf{y}\|_{1-\text{var}} \max(S, R) \right)^{m}}{(m!)^{2}} \quad \text{for all } k \in [\tilde{d}], \end{aligned}$$
(53)

where in (a) we used Corollary C.7, in (b) we used (53). We are almost ready to use Theorem A.13, but it requires the $\|\cdot\|_{\Psi_{1/m}}$ bound in terms of centered random variables. Although $\|\cdot\|_{\Psi_{\alpha}}$ does not satisfy the triangle inequality for $\alpha \in (0, 1)$, it obeys that (see [30, Lemma A.3.]) $\|X + Y\|_{\Psi_{\alpha}} \leq 2^{1/\alpha} (\|X\|_{\Psi_{\alpha}} + \|Y\|_{\Psi_{\alpha}})$ for any random variables *X* and *Y*. For a constant $c \in \mathbb{R}$, we have $\|c\|_{\Psi_{1/m}} = \frac{|c|}{\ln^{m} 2}$, and hence by Lemma C.5 we have that

$$\left\|\mathsf{k}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y})\right\|_{\Psi_{1/m}} \leq \frac{\left(L^{2} \|\mathbf{x}\|_{1-\mathrm{var}} \|\mathbf{y}\|_{1-\mathrm{var}} / \ln 2\right)^{m}}{(m!)^{2}},$$

where $L = \|\mathbb{E}_{\mathbf{w} \sim \Lambda}[\mathbf{w}\mathbf{w}^{\top}]\|_2$ is the Lipschitz constant of the kernel $\mathbf{k} : \mathcal{X} \times \mathcal{X} \to \mathbb{R}$. This gives

$$\begin{aligned} \left\| \hat{\mathsf{k}}_{\text{Sig}_{m}}^{(k)}(\mathbf{x},\mathbf{y}) - \mathsf{k}_{\text{Sig}_{m}}(\mathbf{x},\mathbf{y}) \right\|_{\Psi_{1/m}} &\leq 2^{m} \left(\left\| \hat{\mathsf{k}}_{\text{Sig}_{m}}^{(k)}(\mathbf{x},\mathbf{y}) \right\|_{\Psi_{1/m}} + \left\| \mathsf{k}_{\text{Sig}_{m}}(\mathbf{x},\mathbf{y}) \right\|_{\Psi_{1/m}} \right) \\ &\leq \frac{\left(2 \left\| \mathbf{x} \right\|_{1-\text{var}} \left\| \mathbf{y} \right\|_{1-\text{var}} \right)^{m}}{(m!)^{2}} \left(\left(2d \max(S,R) \right)^{m} + \left(L^{2}/\ln 2 \right)^{m} \right). \end{aligned}$$

Finally, we have the required Orlicz norm bound for invoking Theorem A.13, so that we get

$$\mathbb{P}\left[\left|\tilde{\mathsf{k}}_{\text{Sig}_{m}}^{\text{DP}}(\mathbf{x},\mathbf{y})-\mathsf{k}_{\text{Sig}_{m}}(\mathbf{x},\mathbf{y})\right| \geq \epsilon\right]$$

$$\leq 2\exp\left(-\frac{1}{4}\min\left\{\left(\frac{\sqrt{\tilde{d}}\epsilon}{2C_{d,m}}\right)^{2},\left(\frac{\tilde{d}\epsilon}{\sqrt{8}C_{d,m}}\right)^{\frac{1}{m}}\right\}\right),$$

where the constant $C_{d,m} > 0$ is defined as

$$C_{d,m} \coloneqq \sqrt{8}e^{4}(2\pi)^{1/4}e^{1/24}\frac{(4e\,m\,\|\mathbf{x}\|_{1-\mathrm{var}}\,\|\mathbf{y}\|_{1-\mathrm{var}})^{m}}{(m!)^{2}}\left((2d\,\max(S,R))^{m} + \left(L^{2}/\ln 2\right)^{m}\right),$$

and invoking Stirling's approximation $\frac{1}{m!} \le \left(\frac{e}{m}\right)^m$ gives the stated result.

Now, we prove the analogous result for $\tilde{k}_{Sig_m}^{TRP}$.

Theorem C.12 (Concentration inequality for RFSF-TRP kernel). Let $k : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}$ be a continuous, bounded, translation-invariant kernel with spectral measure Λ , which satisfies for some S, R > 0 that

$$\mathbb{E}_{\mathbf{w}\sim\Lambda}\left[|w_i|^{2k}\right] \leq \frac{k!S^2R^{k-2}}{2} \quad for all \ i \in [d] \ and \ k \geq 2.$$
(54)

Then, for the level-m RFSF-TRP kernel as defined in (21), we have for $\mathbf{x}, \mathbf{y} \in \mathcal{X}_{seq}$ and $\epsilon > 0$

$$\mathbb{P}\left[\left|\tilde{\mathsf{K}}_{\mathrm{Sig}_{m}}^{\mathrm{TRP}}(\mathbf{x},\mathbf{y})-\tilde{\mathsf{K}}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y})\right|\geq\epsilon\right]\leq C_{d,\Lambda}\exp\left(-\left(\frac{m^{2}\tilde{d}_{\mathrm{TRP}}^{\frac{1}{2m}}\epsilon^{\frac{1}{m}}}{2\sqrt{2}e^{3}R\|\mathbf{x}\|_{1}}\right)^{\frac{1}{2}}\right),$$

where $C_{d,\Lambda} \coloneqq 2\left(1 + \frac{S}{2R} + \frac{S^2}{4R^2}\right)^d$.

Proof. First, we consider the conditional probability $\mathbb{P}\left[\left|\tilde{k}_{\text{Sig}_{m}}^{\text{TRP}}(\mathbf{x}, \mathbf{y}) - \tilde{k}_{\text{Sig}_{m}}(\mathbf{x}, \mathbf{y})\right| \ge \epsilon \left|\mathbf{W}\right]$ by conditioning on the RFSF weights $\mathbf{W} \coloneqq (\mathbf{W}^{(1)}, \dots, \mathbf{W}^{(m)})$, so that the only source of randomness comes from the TRP weights $\mathbf{P} \coloneqq (\mathbf{P}^{(1)}, \dots, \mathbf{P}^{(m)})$. The idea is to call Theorem A.15 to estimate the conditional probability, and then take expectation over \mathbf{W} . Since Theorem A.15 quantifies the concentration of a Gaussian polynomial around its mean in terms of its variance, we first compute the conditional statistics of $\tilde{k}_{\text{Sig}_{m}}^{\text{TRP}}(\mathbf{x}, \mathbf{y})$.

Conditional expectation. Recall the definition of $\tilde{\varphi}_{\text{Sig}_m}^{\text{TRP}}(\mathbf{x})$ (20), where $\mathbf{p}_1^{(1)}, \dots, \mathbf{p}_{\tilde{d}_{\text{TRP}}}^{(m)} \overset{\text{i.i.d.}}{\sim} \mathcal{N}(0, \mathbf{I}_{2\tilde{d}})$, and $\mathbf{x} \in \mathcal{X}_{\text{seq}}$, so that

$$\tilde{\varphi}_{\text{Sig}_{m}}^{\text{TRP}}(\mathbf{x}) = \frac{1}{\sqrt{\tilde{d}_{\text{TRP}}}} \left(\sum_{\mathbf{i} \in \Delta_{m}(\ell_{\mathbf{x}})} \prod_{p=1}^{m} \left\langle \mathbf{p}_{i}^{(p)}, \delta \, \tilde{\varphi}_{p}(\mathbf{x}_{i_{p}}) \right\rangle \right)_{i=1}^{d_{\text{TRP}}}.$$

Hence, $\tilde{k}_{\text{Sig}_m}^{\text{TRP}}(\mathbf{x}, \mathbf{y})$ is written for $\mathbf{x}, \mathbf{y} \in \mathcal{X}_{\text{seq}}$ as

$$\tilde{\mathsf{K}}_{\mathrm{Sig}_{m}}^{\mathrm{TRP}}(\mathbf{x},\mathbf{y}) = \frac{1}{\tilde{d}_{\mathrm{TRP}}} \sum_{i=1}^{d_{\mathrm{TRP}}} \sum_{\substack{i=\Delta_{m}(\ell_{\mathbf{x}}) \\ \mathbf{y} \in \Delta_{m}(\ell_{\mathbf{y}})}} \prod_{p=1}^{m} \left\langle \mathbf{p}_{i}^{(p)}, \delta \tilde{\varphi}_{p}(\mathbf{x}_{i_{p}}) \right\rangle \left\langle \mathbf{p}_{i}^{(p)}, \delta \tilde{\varphi}_{p}(\mathbf{y}_{i_{p}}) \right\rangle,$$

which is a sample average of \tilde{d}_{TRP} i.i.d. terms, i.e. $\tilde{k}_{Sig_m}^{TRP}(\mathbf{x}, \mathbf{y}) = \frac{1}{\tilde{d}_{TRP}} \sum_{i=1}^{\tilde{d}_{TRP}} A_i$. We only have to verify that A_i is conditionally an unbiased approximator of \tilde{k}_{Sig_m} given **W**.

$$\mathbb{E}[A_i | \mathbf{W}] \stackrel{\text{(a)}}{=} \sum_{\substack{\mathbf{i} \in \Delta_m(\ell_{\mathbf{x}}) \\ \mathbf{j} \in \Delta_m(\ell_{\mathbf{y}})}} \prod_{p=1}^m \mathbb{E}\left[\left\langle \mathbf{p}_i^{(p)}, \delta \tilde{\varphi}_p(\mathbf{x}_{i_p}) \right\rangle \left\langle \mathbf{p}_i^{(p)}, \delta \tilde{\varphi}_p(\mathbf{y}_{j_p}) \right\rangle \right| \mathbf{W} \right]$$

$$\stackrel{\text{(b)}}{=} \sum_{\substack{\mathbf{i} \in \Delta_m(\ell_{\mathbf{x}}) \\ \mathbf{j} \in \Delta_m(\ell_{\mathbf{y}})}} \prod_{p=1}^m \left\langle \mathbb{E} \left[\mathbf{p}_i^{(p)} \otimes \mathbf{p}_i^{(p)} \right], \delta \tilde{\varphi}_p(\mathbf{x}_{i_p}) \otimes \delta \tilde{\varphi}_p(\mathbf{y}_{j_p}) \right\rangle$$

$$\stackrel{\text{(c)}}{=} \sum_{\substack{\mathbf{i} \in \Delta_m(\ell_{\mathbf{x}}) \\ \mathbf{j} \in \Delta_m(\ell_{\mathbf{y}})}} \prod_{p=1}^m \left\langle I_{\tilde{d}}, \delta \tilde{\varphi}_p(\mathbf{x}_{i_p}) \otimes \delta \tilde{\varphi}_p(\mathbf{y}_{j_p}) \right\rangle$$

$$\stackrel{\text{(d)}}{=} \sum_{\substack{\mathbf{i} \in \Delta_m(\ell_{\mathbf{x}}) \\ \mathbf{j} \in \Delta_m(\ell_{\mathbf{y}})}} \prod_{p=1}^m \left\langle \delta \tilde{\varphi}_p(\mathbf{x}_{i_p}), \delta \tilde{\varphi}_p(\mathbf{y}_{j_p}) \right\rangle,$$

where (a) follows from linearity of expectation and independence of the $\mathbf{p}_i^{(p)}$'s for $p \in [m]$, (b) from bilinearity of inner product, and independence of \mathbf{P} and \mathbf{W} , (c) from substituting the covariance, (d) is since the outer product is projected onto the diagonal.

Conditional variance. We compute the conditional variance of A_i given **W**:

$$\begin{split} \mathbb{E}\left[A_{m,i}^{2}\left|\mathbf{W}\right] \\ &\stackrel{(9)}{=} \sum_{\substack{\mathbf{i},\mathbf{k}\in\Delta_{m}(|\mathbf{x}|) \ p=1}} \mathbb{E}\left[\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{x}_{i_{p}})\right\rangle\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{y}_{j_{p}})\right\rangle\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{x}_{k_{p}})\right\rangle\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{y}_{l_{p}})\right\rangle\left|\mathbf{W}\right] \\ &\stackrel{(1)}{=} \sum_{\substack{\mathbf{i},\mathbf{k}\in\Delta_{m}(|\mathbf{x}|) \ p=1}} \frac{1}{\left(\mathbb{E}\left[\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{x}_{i_{p}})\right\rangle\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{x}_{k_{p}})\right\rangle\left|\mathbf{W}\right]\mathbb{E}\left[\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{x}_{k_{p}})\right\rangle\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{y}_{k_{p}})\right\rangle\left|\mathbf{W}\right] \\ &+ \mathbb{E}\left[\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{x}_{i_{p}})\right\rangle\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{x}_{k_{p}})\right\rangle\right|\mathbf{W}\right]\mathbb{E}\left[\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{y}_{j_{p}})\right\rangle\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{y}_{k_{p}})\right\rangle\left|\mathbf{W}\right] \\ &+ \mathbb{E}\left[\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{x}_{i_{p}})\right\rangle\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{y}_{k_{p}})\right\rangle\right|\mathbf{W}\right]\mathbb{E}\left[\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{x}_{k_{p}})\right\rangle\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{y}_{j_{p}})\right\rangle\left|\mathbf{W}\right]\right] \\ &+ \mathbb{E}\left[\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{x}_{i_{p}})\right\rangle\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{y}_{i_{p}})\right\rangle\right|\mathbf{W}\right]\mathbb{E}\left[\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{x}_{k_{p}})\right\rangle\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{y}_{i_{p}})\right\rangle\left|\mathbf{W}\right]\right] \\ &+ \mathbb{E}\left[\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{x}_{i_{p}})\right\rangle\left\langle\mathbf{p}_{i}^{(p)},\delta\bar{\varphi}_{p}(\mathbf{y}_{i_{p}})\right\rangle\right|\mathbf{W}\right] \\ &+ \mathbb{E}\left[\left\langle\bar{\varphi}_{p}(\mathbf{x}_{i_{p}}),\delta\bar{\varphi}_{p}(\mathbf{y}_{j_{p}})\right\rangle\left\langle\delta\bar{\varphi}_{p}(\mathbf{x}_{k_{p}}),\delta\bar{\varphi}_{p}(\mathbf{y}_{i_{p}})\right\rangle\right] \\ &+ \mathbb{E}\left[\left\langle\bar{\varphi}_{p}(\mathbf{x}_{i_{p}}),\delta\bar{\varphi}_{p}(\mathbf{y}_{i_{p}})\right\rangle\left\langle\delta\bar{\varphi}_{p}(\mathbf{x}_{k_{p}}),\delta\bar{\varphi}_{p}(\mathbf{y}_{i_{p}})\right\rangle\right] \\ &+ \left\langle\bar{\varphi}\bar{\varphi}_{p}(\mathbf{x}_{i_{p}}),\delta\bar{\varphi}_{p}(\mathbf{y}_{i_{p}})\right\rangle\left\langle\bar{\varphi}\bar{\varphi}_{p}(\mathbf{x}_{k_{p}}),\delta\bar{\varphi}_{p}(\mathbf{y}_{i_{p}})\right\right\| \\ &+ \left\langle\bar{\varphi}\bar{\varphi}_{p}(\mathbf{x}_{i_{p}})\right\rangle\left\langle\bar{\varphi}\bar{\varphi}_{p}(\mathbf{x}_{k_{p}})\right\rangle\left\langle\bar{\varphi}\bar{\varphi}_{p}(\mathbf{x}_{k_{p}})\right\rangle\left\langle\bar{\varphi}\bar{\varphi}_{p}(\mathbf{y}_{i_{p}})\right\right\| \\ &+ \left\langle\bar{\varphi}\bar{\varphi}_{p}(\mathbf{x}_{i_{p}})\right\rangle\left\langle\bar{\varphi}\bar{\varphi}_{p}(\mathbf{y}_{i_{p}})\right\|\left\|\delta\bar{\varphi}\bar{\varphi}_{p}(\mathbf{y}_{i_{p}})\right\| \\ &+ \left\langle\bar{\varphi}\bar{\varphi}_{p}(\mathbf{x}_{i_{p}})\right\rangle\left\langle\bar{\varphi}\bar{\varphi}_{p}(\mathbf{x}_{k_{p}})\right\rangle\left\langle\bar{\varphi}\bar{\varphi}\bar{\varphi}_{p}(\mathbf{x}_{k_{p}})\right\rangle\left\langle\bar{\varphi}\bar{\varphi}\bar{\varphi}_{p}(\mathbf{y}_{i_{p}})\right\right\| \\ &= \left\langle\mathbf{x}_{i_{p}}\left(\mathbf{x}_{i_{p}}\right)\left\langle\bar{\varphi}\bar{\varphi}_{p}(\mathbf{x}_{i_{p}})\right\right)\left\langle\bar{\varphi}\bar{\varphi}_{p}(\mathbf{x}_{i_{p}})\right\rangle\right\| \\ &= \left\langle\mathbf{x}_$$

where (e) follows from linearity of expectation and independence of the $\mathbf{p}_i^{(p)}$'s for $p \in [m]$, (f) from Isserlis' theorem [34], (g) is the same as (a)-(d), (h) is the Cauchy-Schwarz inequality, (i) from factorizing the summation, (j) is the same as Lemma C.6.

Therefore, we have due to Lemma C.7 for the variance that

$$\mathbb{V}[A_{m,i} | \mathbf{W}] = \mathbb{E}[A_{m,i}^{2} | \mathbf{W}] - \mathbb{E}^{2}[A_{m,i} | \mathbf{W}] \leq \frac{3^{m} + 1}{(m!)^{4}} \left(\frac{\|\mathbf{x}\|_{1-\operatorname{var}}^{2} \|\mathbf{y}\|_{1-\operatorname{var}}^{2}}{\tilde{d}^{2}}\right)^{m} \prod_{p=1}^{m} \|\mathbf{W}^{(p)}\|_{2}^{4}.$$

Let $\beta_m(\mathbf{x}, \mathbf{y}) \coloneqq \frac{3^m + 1}{(m!)^4} \|\mathbf{x}\|_{1\text{-var}}^{2m} \|\mathbf{y}\|_{1\text{-var}}^{2m}$. Then, as $\tilde{k}_{\text{Sig}_m}^{\text{TRP}}(\mathbf{x}, \mathbf{y}) \|\mathbf{W}$ is a sample average,

$$\mathbb{V}\left[\tilde{\mathsf{K}}_{\text{Sig}_{m}}^{\text{TRP}}(\mathbf{x},\mathbf{y}) \middle| \mathbf{W}\right] \leq \frac{\beta_{m}(\mathbf{x},\mathbf{y})}{\tilde{d}_{\text{TRP}}\tilde{d}^{2m}} \prod_{p=1}^{m} \left\| \mathbf{W}^{(p)} \right\|_{2}^{4}.$$
(55)

Conditional bound. Since $\tilde{k}_{\text{Sig}_m}^{\text{TRP}}(\mathbf{x}, \mathbf{y}) | \mathbf{W}$ is a Gaussian polynomial of degree-2*m*, with expectation $\tilde{k}_{\text{Sig}_m}(\mathbf{x}, \mathbf{y})$, and variance (55), we have by Theorem A.15 for $\epsilon > 0$ that

$$\mathbb{P}\left[\left|\tilde{\mathsf{k}}_{\mathrm{Sig}_{m}}^{\mathrm{TRP}}(\mathbf{x},\mathbf{y}) - \tilde{\mathsf{k}}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y})\right| \ge \epsilon \left|\mathbf{W}\right] \le 2 \exp\left(-\frac{\epsilon^{\frac{1}{m}}}{2\sqrt{2}e^{\bigvee \frac{1}{2m}}\left[\tilde{\mathsf{k}}_{\mathrm{Sig}_{m}}^{\mathrm{TRP}}(\mathbf{x},\mathbf{y})\right]\mathbf{W}\right]}\right)$$
$$\le 2 \exp\left(-\frac{\tilde{d}_{\mathrm{TRP}}^{\frac{1}{2m}}\tilde{d}\epsilon^{\frac{1}{m}}}{2\sqrt{2}e^{\beta_{m}^{\frac{1}{2m}}}(\mathbf{x},\mathbf{y})\prod_{p=1}^{m}\left\|\mathbf{W}^{(p)}\right\|_{2}^{\frac{2}{m}}}}\right).$$
(56)

Undoing the conditioning. We take the expectation in (56) so that

$$\mathbb{P}\left[\left|\tilde{\mathsf{K}}_{\mathrm{Sig}_{m}}^{\mathrm{TRP}}(\mathbf{x},\mathbf{y})-\tilde{\mathsf{K}}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y})\right| \geq \epsilon\right] = \mathbb{E}\left[\mathbb{P}\left[\left|\tilde{\mathsf{K}}_{\mathrm{Sig}_{m}}^{\mathrm{TRP}}(\mathbf{x},\mathbf{y})-\tilde{\mathsf{K}}_{\mathrm{Sig}_{m}}(\mathbf{x},\mathbf{y})\right| \geq \epsilon\left|\mathbf{W}\right]\right] \\ \leq 2\mathbb{E}\left[\exp\left(-\frac{d\tilde{t}_{\mathrm{TRP}}^{\frac{1}{2m}}d\epsilon\frac{1}{m}}{2\sqrt{2}e\beta_{m}(\mathbf{x},\mathbf{y})^{\frac{1}{2m}}}\prod_{p=1}^{m}\left\|\mathbf{W}^{(p)}\right\|_{2}^{\frac{2}{m}}}\right)\right] \\ \stackrel{(k)}{\leq} 2\mathbb{E}\left[\exp\left(-\frac{\lambda d\tilde{t}_{\mathrm{TRP}}^{\frac{1}{2m}}\epsilon^{\frac{1}{m}}}{2\sqrt{2}e\beta_{m}(\mathbf{x},\mathbf{y})^{\frac{1}{2m}}}\frac{d\tilde{d}}{\lambda\prod_{p=1}^{m}\left\|\mathbf{W}^{(p)}\right\|_{2}^{\frac{2}{m}}}\right)\right] \\ \stackrel{(l)}{\leq} 2\mathbb{E}\left[\exp\left(-2\left(\frac{\lambda d\tilde{t}_{\mathrm{TRP}}^{\frac{1}{2m}}\epsilon^{\frac{1}{m}}}{2\sqrt{2}e\beta_{m}(\mathbf{x},\mathbf{y})^{\frac{1}{2m}}}\right)^{\frac{1}{2}} + \frac{\lambda\prod_{p=1}^{m}\left\|\mathbf{W}^{(p)}\right\|_{2}^{\frac{2}{m}}}{d}\right)\right] \\ \stackrel{(m)}{\leq} 2\mathbb{E}\left[\exp\left(-\left(\frac{\sqrt{2}\lambda d\tilde{t}_{\mathrm{TRP}}^{\frac{1}{2m}}\epsilon^{\frac{1}{m}}}{e\beta_{m}(\mathbf{x},\mathbf{y})^{\frac{1}{2m}}}\right)^{\frac{1}{2}} + \frac{\lambda\sum_{p=1}^{m}\left\|\mathbf{W}^{(p)}\right\|_{2}^{2}}{m\tilde{d}}\right)\right], \quad (57)$$

where (k) follows form multiplying and dividing with a $\lambda > 0$, (l) from applying Lemma A.4 with $p = \frac{1}{2}$ and q = 1, and (m) from the arithmetic-geometric mean inequality. Bounding the MGF of $w_{i,j}^{(p)^2}$ for $p \in [m], i \in [d], j \in [\tilde{d}]$, we have that

$$\mathbb{E}\left[\exp\left(\lambda w_{i,j}^{(p)^2}\right)\right] \stackrel{(n)}{\leq} \sum_{k\geq 0} \mathbb{E}\left[w_{i,j}^{(p)^k}\right] \frac{\lambda^k}{k!} \stackrel{(o)}{\leq} 1 + \lambda S + \frac{\lambda^2 S^2}{2} \sum_{k\geq 0} (\lambda R)^k$$
$$\stackrel{(p)}{=} 1 + \lambda S + \frac{\lambda^2 S^2}{2} \frac{1}{1 - \lambda R} \stackrel{(q)}{=} 1 + \frac{S}{2R} + \frac{S^2}{4R^2},$$

where (n) is the Taylor expansion, (o) is the condition (54) and applying Jensen inequality to the degree-1 term, (p) is the geometric series for $\lambda < \frac{1}{R}$, and (q) is choosing $\lambda = \frac{1}{2R}$. Hence,

$$\mathbb{E}\left[\exp\left(\frac{\lambda\sum_{p=1}^{m}\|\mathbf{W}^{(p)}\|_{2}^{2}}{m\tilde{d}}\right)\right] = \mathbb{E}\left[\exp\left(\frac{\lambda\sum_{p=1}^{m}\sum_{i=1}^{d}\sum_{j=1}^{\tilde{d}}w_{i,j}^{(p)^{2}}}{m\tilde{d}}\right)\right]$$
$$\stackrel{(r)}{\leq} \mathbb{E}^{1/(m\tilde{d})}\left[\exp\left(\lambda\sum_{p=1}^{m}\sum_{i=1}^{d}\sum_{j=1}^{\tilde{d}}w_{i,j}^{(p)^{2}}\right)\right]$$
$$\stackrel{(s)}{\leq} \left(1 + \frac{S}{2R} + \frac{S^{2}}{4R^{2}}\right)^{d},$$

where (r) is due to the Jensen inequality (Lemma A.1), and (s) follows from the independence of the $w_{i,j}^{(p)}$'s for $p \in [m], i \in [d], j \in [\tilde{d}]$. Finally, plugging this into (57), we get that

$$\mathbb{P}\left[\left|\tilde{\mathsf{K}}_{\operatorname{Sig}_{m}}^{\operatorname{TRP}}(\mathbf{x},\mathbf{y}) - \tilde{\mathsf{K}}_{\operatorname{Sig}_{m}}(\mathbf{x},\mathbf{y})\right| \ge \epsilon\right] \le 2\left(1 + \frac{S}{2R} + \frac{S^{2}}{4R^{2}}\right)^{d} \exp\left(-\left(\frac{\tilde{d}_{\operatorname{TRP}}^{\frac{1}{2m}} \epsilon^{\frac{1}{m}}}{\sqrt{2}eR\beta_{m}(\mathbf{x},\mathbf{y})^{\frac{1}{2m}}}\right)^{\frac{1}{2}}\right).$$

Finally, note that $\beta_m(\mathbf{x}, \mathbf{y})^{\frac{1}{2m}} = \left(\frac{3^m+1}{(m!)^4}\right)^{\frac{1}{2m}} \|\mathbf{x}\|_{1-\operatorname{var}} \|\mathbf{y}\|_{1-\operatorname{var}} \le \frac{2e^2}{m^2} \|\mathbf{x}\|_{1-\operatorname{var}} \|\mathbf{y}\|_{1-\operatorname{var}}$, since $3^m + 1 \le 4^m$ for $m \ge 1$, and $\frac{1}{m!} \le \left(\frac{e}{m}\right)^m$ due to Stirling's approximation.

D Algorithms

We adopt the following notation for describing vectorized algorithms from [39, 82]. For arrays, 1-based indexing is used. Let *A* and *B* be *k*-fold arrays with shape $(n_1 \times \cdots \times n_k)$, and let $i_j \in [n_j]$ for $j \in [k]$. We define the following array operations:

(i) The cumulative sum along axis *j*:

$$A[\ldots,:,\boxplus,:][\ldots,i_{j-1},i_j,,i_{j+1}\ldots] := \sum_{\kappa=1}^{i_j} A[\ldots,i_{j-1},\kappa,i_{j+1},\ldots].$$

(ii) The slice-wise sum along axis *j*:

$$A[\ldots, :, \Sigma, :, \ldots][\ldots, i_{j-1}, i_{j+1}, \ldots] := \sum_{\kappa=1}^{n_j} A[\ldots, i_{j-1}, \kappa, i_{j+1}, \ldots].$$

(iii) The shift along axis *j* by +m for $m \in \mathbb{Z}+$:

$$A[\dots, :, +m, :, \dots][\dots, i_{j-1}, i_j, i_{j+1}, \dots]$$

:=
$$\begin{cases} A[\dots, i_{j-1}, i_j - m, i_{j+1}, \dots], & \text{if } i_j > m \\ 0 & \text{if } i_j \le m \end{cases}$$

(iv) The Hadamard product of arrays A and B:

$$(A \odot B)[i_1,\ldots,i_k] := A[i_1,\ldots,i_k]B[i_1,\ldots,i_k].$$

(v) Now, if *A* has shape $(n_1 \times \cdots \times n_j \times \cdots \times n_k)$ and *B* has shape $(n_1 \times \cdots \times n'_j \times \cdots \times n_k)$, then their (batch) outer product along axis j is defined for $i_i \in [n_i n'_i]$ as

 $(A \boxtimes_i B)[i_1, \ldots, i_j, \ldots, i_k] \coloneqq A[i_1, \cdots, [i_i/n'_i], \cdots, i_k]B[i_1, \ldots, i_j \mod n_j, \ldots, i_k],$

where $[\cdot] : \mathbb{R} \to \mathbb{Z}$ refers to the ceiling operation, and $(\cdot \mod n) : \mathbb{Z} \to [n]$ to the modulo *n* operation that maps onto [n] for $n \in \mathbb{Z}$.

Algorithm 1 Computing the RFSF map $\tilde{\varphi}_{Sig_{< M}}$.

1: Input: Sequences $\mathbf{X} = (\mathbf{x}_i)_{i=1}^N \subset \mathcal{X}_{seq}$, measure Λ , truncation $M \in \mathbb{Z}_+$, RFF sample size $\tilde{d} \in \mathbb{Z}_+$

- 2: Optional: Add time-parametrization $\mathbf{x}_i \leftarrow (\mathbf{x}_{i,t}, t/\ell_{\mathbf{x}_i})_{t=1}^{\ell_{\mathbf{x}_i}}$ for all $i \in [N]$
- 3: Tabulate to uniform length $\ell \coloneqq \max_{j \in [N]} \ell_{\mathbf{x}_j}$ by $\mathbf{x}_i \leftarrow (\mathbf{x}_{i,1}, \dots, \mathbf{x}_{i,\ell_{\mathbf{x}_i}}, \dots, \mathbf{x}_{i,\ell_{\mathbf{x}_i}})$ for all $i \in [N]$
- 4: Sample independent RFF weights $\mathbf{W}^{(1)}, \dots, \mathbf{W}^{(M)} \stackrel{\text{i.i.d.}}{\sim} \Lambda^{\tilde{d}}$
- 5: Initialize an array U with shape $[M, N, L-1, 2\tilde{d}]$
- 6: Compute increments $U[m, i, t, :] \leftarrow \tilde{\varphi}_m(\mathbf{x}_{i,t+1}) \tilde{\varphi}_m(\mathbf{x}_{i,t})$ for $m \in [M], t \in [L-1]$
- 7: Initialize array $V \leftarrow U[1, :, :, :]$
- 8: Collapse into level-1 features $P_1 \leftarrow V[:, \Sigma, :]$
- 9: **for** m = 2 **to** *M* **do**
- 10: Update with next increment $V \leftarrow V[:, \boxplus +1, :] \boxtimes_3 U[m, :, :, :]$
- 11: Collapse into level-*m* features $P_m \leftarrow V[:, \Sigma, :]$
- 12: end for
- 13: **Output:** Arrays of RFSF features per signature level P_1, \ldots, P_M .

Algorithm 2 Computing the RFSF-DP map $\tilde{\varphi}_{\text{Sig}_{\leq M}}^{\text{DP}}$.

1: **Input:** Sequences $\mathbf{X} = (\mathbf{x}_i)_{i=1}^N \subset \mathcal{X}_{seq}$, measure Λ , truncation $M \in \mathbb{Z}_+$, RFSF-DP sample size $\tilde{d} \in \mathbb{Z}_+$

- 2: Optional: Add time-parametrization $\mathbf{x}_i \leftarrow (\mathbf{x}_{i,t}, t/\ell_{\mathbf{x}_i})_{t=1}^{\ell_{\mathbf{x}_i}}$ for all $i \in [N]$ 3: Tabulate to uniform length $\ell \coloneqq \max_{j \in [N]} \ell_{\mathbf{x}_j}$ by $\mathbf{x}_i \leftarrow (\mathbf{x}_{i,1}, \dots, \mathbf{x}_{i,\ell_{\mathbf{x}_i}}, \dots, \mathbf{x}_{i,\ell_{\mathbf{x}_i}})$ for all $i \in [N]$
- 4: Sample independent RFF weights $\mathbf{W}^{(1)}, \dots, \mathbf{W}^{(M)} \stackrel{\text{i.i.d.}}{\sim} \Lambda^{\tilde{d}}$
- 5: Initialize an array U with shape $[M, N, L-1, \tilde{d}, 2]$
- 6: Compute increments $U[m, i, t, k, :] \leftarrow \hat{\varphi}_{m,k}(\mathbf{x}_{i,t+1}) \hat{\varphi}_{m,k}(\mathbf{x}_{i,t})$ for $m \in [M]$, $i \in [N]$, $t \in [L-1]$, $k \in [\tilde{d}]$ 7: Initialize array $V \leftarrow \frac{1}{\sqrt{d}}U[1, :, :, :]$ 8: Collapse into level-1 features $P_1 \leftarrow V[:, \Sigma, :, :]$

- 9: **for** m = 2 **to** *M* **do**
- Update with next increment $V \leftarrow V[:, \boxplus +1, :] \boxtimes_4 U[m, :, :, :]$ 10:
- Collapse into level-*m* features $P_m \leftarrow V[:, \Sigma, :, :]$ 11:
- 12: end for
- 13: **Output:** Arrays of RFSF-DP features per signature level P_1, \ldots, P_M .

Algorithm 3 Computing the RFSF-TRP map $\tilde{\varphi}_{\text{Sig}_{SM}}^{\text{TRP}}$.

1: **Input:** Sequences $\mathbf{X} = (\mathbf{x}_i)_{i=1}^N \subset \mathcal{X}_{seq}$, measure Λ , truncation $M \in \mathbb{Z}_+$, RFSF and TRP sample size $\tilde{d} \in \mathbb{Z}_+$

2: Optional: Add time-parametrization $\mathbf{x}_i \leftarrow (\mathbf{x}_{i,t}, t/\ell_{\mathbf{x}_i})_{t=1}^{\ell_{\mathbf{x}_i}}$ for all $i \in [N]$ 3: Tabulate to uniform length $\ell \coloneqq \max_{j \in [N]} \ell_{\mathbf{x}_j}$ by $\mathbf{x}_i \leftarrow (\mathbf{x}_{i,1}, \dots, \mathbf{x}_{i,\ell_{\mathbf{x}_i}}, \dots, \mathbf{x}_{i,\ell_{\mathbf{x}_i}})$ for all $i \in [N]$

4: Sample independent RFF weights $\mathbf{W}^{(1)}, \dots, \mathbf{W}^{(M)} \stackrel{\text{i.i.d.}}{\sim} \Lambda^{\tilde{d}}$

5: Sample standard normal matrices $\mathbf{P}^{(1)}, \dots, \mathbf{P}^{(M)} \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}^{\tilde{d}}(\mathbf{0}, \mathbf{I}_{2\tilde{d}})$

6: Initialize an array *U* with shape $[M, N, L-1, \tilde{d}]$

7: Compute projected increments $U[m, i, t, :] \leftarrow \mathbf{P}^{(m)^{\top}} \left(\tilde{\varphi}_m(\mathbf{x}_{i,t+1}) - \tilde{\varphi}_m(\mathbf{x}_{i,t}) \right)$ for $m \in [M]$, $i \in [N]$, $t \in [L-1]$ 8: Initialize array $V \leftarrow \frac{1}{\sqrt{d}} U[1, :, :]$ 9: Collapse into level-1 features $P_1 \leftarrow V[:, \Sigma, :]$

10: **for** m = 2 **to** M **do**

Update with next increment $V \leftarrow V[:,\boxplus+1,:] \odot U[m,:,:,:]$ 11:

12: Collapse into level-*m* features $P_m \leftarrow V[:, \Sigma, :]$

13: end for

14: **Output:** Arrays of RFSF-TRP features per signature level P_1, \ldots, P_M .