

Multivariate Morphometrics and Cytotaxonomy
of the West African
Simulium damnosum Complex (Diptera: Simuliidae).

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By

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ABSTRACT

The purpose of this project was to examine variation within the West African *Simulium damnosum* species complex using classical larval polytene chromosome analysis, multivariate statistical analysis of larval polytene chromosome variation, and multivariate statistical analysis of adult female morphological variation.

Classical polytene chromosome analysis, undertaken to provide correlated identification for adult females reared from pupae in the same samples, revealed a chromosomally distinct form within *S. sanctipauli* from Togo. This form was distinguished from typical *S. sanctipauli* by the strong Y-chromosome linkage of an inversion, 1S-21, found autosomally at low frequency in typical *S. sanctipauli*.

Multivariate statistical analysis was applied to available data of larval polytene chromosome inversion frequencies from the *S. sanctipauli* subcomplex, the first application of these methods to polytene chromosome variation. Intra- and Inter-specific variation was found to be more complex than had been found using classical methods of analysis. The most distinctive and internally homogeneous taxon was *S. soubrense* 'B', while *S. sanctipauli* was found to be heterogeneous, with OP-insecticide resistant flies represented as a distinct cluster. *Simulium soubrense* was found to be an heterogeneous assemblage of taxa. The taxa *S. soubrense* 'Chutes Milo' and *S. soubrense* 'Beffa' showed affinities for *S. sanctipauli*, while *S. soubrense* 'Menankaya/Konkoure' showed complex variation which may have been a combination of clinal and local differentiation.

Multivariate morphometric methods were applied to 28 characters of adult females of the West African *S. damnosum* complex. Some of these characters were chosen for their known taxonomic importance, and some as additional characters representing the general morphology.

Statistical methods were undertaken to ensure the integrity of the basic data base, including univariate and multivariate outlier detection procedures.

Multivariate morphometric intraspecific variation was analysed in the six main species of the *S. damnosum* complex, and shape differences correlated with chromosomal differences, seasonal size variation, and both size and shape variation with no clear cause were found in the different species. The predominant mode of variation was found to be size variation, with very strong seasonal size variation for *S. squamosum*. *Simulium soubrense* showed the most extensive morphological variation, although this did not exactly parallel the chromosomal differentiation in the group.

Multivariate interspecific variation was analysed from the perspective of allocatory discriminant analysis, and optimal subsets of characters derived for overall and species-pair analyses from regional and 'global' material.

A method of adjusting prior probabilities of species membership was derived to exploit the taxonomic potential of two non-normal characters, and two kinds of allocation, forced and typicality probability were used to identify flies. The typicality probability method was chosen because it gave approximate confidence intervals to a fly's probability of species membership without reference to the other species. This was the first application of this method to insect morphometrics.

Significant interspecific morphometric variation was found, with most successful identification being of the epidemiologically important species *S. damnosum* s.s., and *S. sirbanum*, and for *S. squamosum*. *Simulium soubrense*, *S. sanctipauli* and *S. yahense* were similar morphologically, although a colour character could identify *S. yahense* with 96% accuracy.

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Appendix Two: Publication, Surtees, D.P., Fiasorgbor, G., Post, R.J.,
Weber, E.A. (1988). The cytotaxonomy of the Djodji form of *Simulium*
sanctipauli (Diptera: Simuliidae). Trop. Med. Parasit., 39, 120-122.
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CHAPTER ONE: INTRODUCTION

1.1 BACKGROUND

The Simuliidae is a widespread family of Nematoceran Diptera, commonly known as blackflies containing about 2000 species-level taxa (Crosskey 1987a). The larvae are aquatic, most requiring running water for filter feeding. In the tropics, larval development can be completed in a few days or weeks, after which the larva enters the pharate or pre-pupal stage, followed by the pupal stage which is spent in a sheltered cocoon spun by the pharate pupa. In the tropics eclosion of the adult often follows within a week. Male blackflies do not bite, but females of most species require a blood meal for egg development. Host species are warm blooded vertebrates, predominantly birds.

A variety of pathogens are transmitted by the females, but the most important human health problem is due to the filarial parasite *Onchocerca volvulus* Leukart, which is transmitted solely by blackflies, and which causes the debilitating disease onchocerciasis. Nearly 18 million people are infected with *O. volvulus* (WHO Technical Report 752, 1987) throughout Tropical Africa, Yemen and parts of Central and South America although this is certainly an underestimate of the true number. By far the major proportion of these people live in sub-saharan West Africa, where members of the *Simulium damnosum* complex are the sole known vectors. *Simulium bovis* De Meillon has been shown to be anthropophilic in northern Nigeria and to support filariae similar to *Onchocerca volvulus*. Its role as a vector of

human onchocerciasis is uncertain, but it is unlikely to be of importance (Crosskey 1957).

In response to this enormous disease problem, the World Health Organisation (WHO) began the Onchocerciasis Control Programme (OCP) in 1974 (Walsh *et al.* 1987), which by 1989 will cover an area of 1 million km² in West Africa. The original aim of OCP was that in the controlled areas annual biting rates ($ABR = \sum(\text{Number of flies caught} \times \text{Number of days in the month}) \div \text{Number of catching days in the month}$, where summation is over the twelve monthly estimates of the monthly biting rate MBR) should be less than 1000, and annual transmission potential ($ATP = \sum(MBR \times \text{Total no. of } L_3 \text{ larvae of } Onchocerca \text{volvulus} \text{ found in the head}) \div \text{No. of flies dissected}$, where summation is over the twelve monthly MTPs) less than 100 for two consecutive years. Up to the present, control of the disease has been effected by larviciding the rivers in which larvae of the *S. damnosum* complex live, with the insecticides temephos, chlorphoxim, permethrin, carbosulfan and *Bacillus thuringiensis* H-14. Future control will include the use of Ivermectin, a microfilaricidal drug, given on a large scale to infected human populations.

The entomological control programme has been successful in the central, well controlled area of the OCP (85% of the original OCP area), in that ABR is usually zero and rarely exceeds 100, whilst ATP is zero throughout the area. However in areas prone to reinvasion of flies by long range migration from uncontrolled areas, in Togo, Benin and Mali, ATPs of up to 1000 have been recorded.

The epidemiological impact of control reflects the success of the entomological control. The community microfilarial load (the geometric mean microfilarial load per skin snip for a cohort of adults)

has decreased linearly by 70% in the central area, but has settled at a higher level than this in areas subject to reinvasion. Ocular onchocerciasis has decreased dramatically in the central area (as measured by the community microfilarial load in the anterior chamber of the eye), though less dramatically in areas subject to reinvasion (WHO technical report 752, 1987).

1.2 THE SIMULIUM DAMNOSUM COMPLEX

Simulium damnosum Theobald (1903) was first implicated as the vector of human onchocerciasis by Blacklock (1926) in Sierra Leone. For the next forty years the species was regarded as a single species. However, over a period of time it became clear that *S. damnosum s.l.* was not uniform throughout its range. Gibbins (1933) noted that adult male thoracic markings varied geographically in East Africa. Grenier and Ovazza (1951) found variation in the fronto-clypeus and in the tubercles of larval *S. damnosum* and Crosskey (1960) also found morphological variation in larval *S. damnosum*. Crisp (1956) measured variation in the size of adult female *S. damnosum* in Northern Ghana and Grenier *et al.* (1960) also found size variation in *S. damnosum*. Lewis (1960) described variation in wing length of adult female *S. damnosum* from Liberia and Cameroon, and also reported variation in proportion of nulliparous flies, biting cycle, midday lull in activity, flight range, egg production, number of flies with retained eggs and degree of man-fly contact in savanna and forest zones. He considered these differences were important enough to influence onchocerciasis transmission. Marr and Lewis (1963, 1964) found variation in the colour of the antennae of adult females from Ghana.

MacCrae (1965, 1967) was the first to consider variation from a taxonomic viewpoint. He examined anthropophilic and non-anthropophilic populations of *S. damnosum s.l.* in Uganda and considered that wing length might be correlated with anthropophilic behaviour. Lewis and Duke (1966) examined morphological and behavioural parameters of *S. damnosum s.l.*. They considered that colour variation was clinal in West Africa, with darker flies inhabiting the forest and lighter flies inhabiting the savanna. They also

examined variation in the tuft of hairs on the fore tarsus, the shape of the fore basitarsus, wing length, the time of biting and the body region of biting. Overall, they thought that the differences between forest and savanna flies was partly due to factors affecting individual variation and partly due to clinal variation.

This variation led Dunbar (1966) to initiate cytotaxonomic studies of East African *S. damnosum s.l.*, using larval silk gland polytene chromosomes, following extensive research on Canadian Simuliidae (Rothfels 1956, 1987, Dunbar 1959). Polytene chromosomes are a special form of polyploid nucleus in which the individual chromosomes are arranged in a highly ordered way with respect to each other (see Ashburner 1979). Because of this ordered arrangement, a consistent pattern of dark bands and light interbands can be read, which is basically constant within species. Structural rearrangements of the chromosome, most commonly paracentric inversions, can be fixed between species. In East Africa Dunbar (1966) recognised four sibling species within the *S. damnosum* complex. A further five were added by Dunbar (1969), with the complex being divided into two subgroups, 'Nile' and 'Sanje'. Dunbar and Vajime (1971, 1972) further divided the complex into 17 species from East and West Africa. In West Africa, Vajime and Dunbar (1975) described eight cytospecies within the *S. damnosum* complex, *S. damnosum s.s.*, *S. sirbanum*, *S. sudanense*, *S. diegeurense*, *S. sanctipauli*, *S. soubrense*, *S. squamosum*, *S. yahense*. The validity of certain of the West African taxa has been questioned (Quillévéré 1975, Quillévéré and Pendriez 1975), but the most recent comprehensive summary (Dunbar and Vajime 1981) reported 26 siblings within the *S. damnosum* complex throughout East and West Africa. Since 1981 the *S. sanctipauli* subcomplex has been revised (Post 1986) with

the addition of a new cytospecies, *S. soubrense* 'B' from Sierra Leone and Guinea and new chromosomal forms continue to be added to the *S. damnosum* complex (e.g Meredith *et al.* 1983, Surtees *et al.* 1988).

The following is a list of the currently recognised taxa within the West African *S. damnosum* complex (Partly based on Crosskey 1987b,

- S. damnosum s.s.*.....savanna, vector
- S. sirbanum*.....savanna, vector
- S. sudanense*.....savanna, vector,
taxonomic status
uncertain (Vajime 1984)
- S. dieguerense*.....savanna, formerly thought
rare, but now considered
more widespread
(Boakye and Mosha 1988),
vector ?
- S. sanctipauli*.....forest, vector
- S. sanctipauli* 'Djodji'.....forest, but may also
extend range into
'savanna
- S. soubrense*.....forest, vector
- S. soubrense* 'Chutes Milo'.....forest, vector
- S. soubrense* 'Konkoure'.....forest, non-vector?
- S. soubrense* 'Beffa'.....forest/savanna, vector
- S. soubrense* 'Menankaya'.....forest, vector
- S. soubrense* 'B'.....forest, vector
- S. squamosum*.....forest/savanna, vector
- S. yahense*.....forest, vector

This list is probably an underestimate of the true number of West African sibling species within *S. damnosum s.l.*, as variation has been noted in some of these taxa, but not fully investigated (Post pers. comm.).

These taxa undoubtedly differ in their importance as vectors of onchocerciasis (Quillévééré 1979) although their exact relative importance has not been fully established because of the problem of identifying adult females of the *S. damnosum* complex, and because of the similar problem of distinguishing between *O. volvulus* and other species of animal *Onchocerca*. However, the basic OCP operational distinction regarding vectorial importance is between savanna dwelling and forest dwelling species (WHO technical report 597, 1976). The savanna dwelling flies (most commonly *S. damnosum s.s.* and *S. sirbanum*) are the most dangerous vectors of onchocerciasis, and control measures have been aimed specifically at controlling these two species, and extension of control has occurred in response to reinvasion of these two species (adult females of these species can migrate distances greater than 500 km, Garms and Walsh 1987) from outside the control areas (Garms *et al.* 1979, Walsh *et al.* 1987).

While the control of onchocerciasis in West Africa is based on the epidemiological and pathological differences between savanna and forest forms of the disease, the identification of adult female savanna flies is not the only important distinction. More complex discrimination between vector species arises in the context of, for example, the identification of insecticide resistant flies (Post and Kurtak 1987), the identification of reinvading flies (e.g. Cheke and Garms 1983), and more detailed local studies of disease transmission (e.g. Garms 1983, in Liberia).

1.3 ADULT IDENTIFICATION METHODS

The major practical motivation for recognising cytotaxa within the *S. damnosum* complex is that they can differ in their capacity to transmit human onchocerciasis (WHO technical report 597, 1977). Cytotaxonomy has mostly been based on larval studies (because only larvae have suitable polytene chromosomes) and a fundamental difficulty in onchocerciasis research remains the inability to distinguish accurately the adult females of most cytospecies within the complex (Phillipon 1987).

Six methods have so far been attempted to identify cytospecies of the *S. damnosum* complex as adult females.

1. Adult Polytene Chromosomes

Procunier (Procunier and Post 1986), building on a technique developed by Bedo (1976) successfully identified adult females of *S. sanctipauli* and *S. soubrense* 'B' caught biting on man. The method uses the same cytotaxonomic criteria (polytene chromosome banding patterns) as was originally used to describe the sibling species within *S. damnosum s.l.*, and is therefore potentially as accurate. The adult polytene chromosomes were taken from the Malpighian tubules. Unfortunately only a low rate of identifiable chromosomes ($\approx 8\%$) could be obtained, and the females needed to be blood-fed (with the consequent ethical problem of feeding potentially infective flies on human volunteers). Wirtz and Raybould (1986) show that artificial blood feeding systems may be practicable for *S. damnosum s.l.* removing one obstacle to the use of the technique, however, unless technical advances can improve the rate of identifiable chromosomes obtained it

is unlikely that the method will be used for routine identification of adult females.

2. Laboratory Reared Larval Progeny.

Raybould *et al.* (1979) reared larvae from wild caught blood fed females, which were then induced to lay eggs in the laboratory. The larvae were then reared in artificial rearing apparatus and identified chromosomally using the chromosome standards of Vajime and Dunbar (1975). The method is therefore as accurate as the chromosomal criteria for identifying flies, but is laborious since it involves the separate rearing of single egg batches. Raybould *et al.* (1979) is significant for showing that all six of the major West African cytospecies were capable of being naturally infective with L_3 s indistinguishable from *O. volvulus*.

3. Enzyme Electrophoresis.

Meredith and Townson (1981) performed an electrophoretic survey of 44 enzyme systems from six species within the West African *S. damnosum* complex from 25 sites in three countries, Mali, Côte d'Ivoire and Ghana. They found that two enzyme systems, phosphoglucomutase (PGM) and trehalase had allozymes which were diagnostic for two species, *S. squamosum* and *S. yahense*. The two species can be distinguished from other members of the *S. damnosum* complex by trehalase A with 98.8% accuracy, and *S. yahense* can be distinguished from *S. squamosum* using PGM B₁ with 99.8% accuracy. These enzymes were used to identify flies caught at human bait as *S. yahense* and *S. squamosum* (Meredith 1982). Garms and Zillman (1984) used the enzyme systems in the field in Liberia to identify *S. yahense* and *S. sanctipauli*,

and found that of the the results which were unequivocal, 99.3% were identified as either *S. sanctipauli* or *S. yahense*, with 0.7% being designated as 'hybrids' (i.e. heterozygotes). They compared these results with a new morphological character found in *S. yahense* (the colour of the setae on the ninth abdominal tergite, see below) to evaluate the taxonomic use of this character. Thomson *et al.* (1988) also compared the results of enzyme electrophoresis with those using morphological characters, and found that the enzyme systems successfully identified *S. yahense* and *S. squamosum*.

Townson *et al.* (1987) recorded geographic variation in allozyme frequencies between *S. squamosum* from Côte d'Ivoire and from Togo. In East Africa Mebrahtu *et al.* (1986) found significant allozymic variation in *S. damnosum s.l.* populations, but the chromosomal identity of these populations was not established.

4. DNA Probes.

The potential importance of using DNA sequences as taxonomic characters is that the genome is being sampled directly, so avoiding any possibility of environmentally mediated variation. Post (1985 and Townson *et al.* 1987) screened about 2000 random sequences from genomic libraries constructed from *S. soubrense* 'B' and *S. squamosum*, and found three DNA sequences which could separate the *S. damnosum* complex into three parts: *S. squamosum/S. yahense*, *S. soubrense* 'B', and *S. sirbanum* according to relative amounts of hybridisation to the three probes using the dot-blot technique. This is an improvement over the results from enzyme electrophoresis because the savanna vectors can be distinguished. However the method is still relatively new and so does not have the backup in practical and the-

oretical understanding from previous experience with other groups of organisms that enzyme electrophoresis enjoys. Also, the samples used for construction of the genomic libraries were collected from a limited geographic area (Sierra Leone), and it appears that there is intraspecific variation in the the probes which prevent their use east of Côte d'Ivoire (Post pers. comm.). Finally, the current use of radioactively labelled probes works against the method as a practical field technique.

5. Analysis of Cuticular Hydrocarbons.

Philips *et al.* (1985 also Townson *et al.* 1987) building on previous work on the *Anopheles gambiae* complex (Carlson and Service 1979,1980) and on the *S. damnosum* complex (Carlson and Walsh 1981) used gas liquid chromatography and gas chromatography/ mass spectrometry to analyse cuticular hydrocarbons of four species within the *S. damnosum* complex, *S. damnosum s.s.*, *S. sirbanum* (two samples), *S. sanctipauli*, and *S. yahense* from four countries in West Africa. Using multivariate analysis of the hydrocarbon profiles from the four species examined, they found that 94.6% of females were reallocated into their correct sample. This included two samples of *S. sirbanum*, so there was significant intraspecific variation in hydrocarbon profiles. It is not clear how many hydrocarbon peaks were used in their discriminant analysis, but if all 24 peaks numbered in figure one of Philips *et al.* (1985) were used then the total sample size of 131 females is too small. Lachenbruch and Goldstein (1979) suggest that the sample size in each group in a discriminant analysis should exceed three times the number of characters i.e. >72 if all 24 peaks were used, otherwise serious bias will be introduced into

the model, making the results misleadingly optimistic. Two of the samples in Philips *et al.* (1985) were very small (*S. yahense*, 9, *S. damnosum s.s.*, 11). These samples are distant from the other three samples in their figure three, which may be due to sampling error.

The statistical problems with the results presented to date may be remedied once larger samples have been obtained, however the method uses expensive equipment requiring technical support, so that it is unlikely that it will be a practical field technique for some time to come.

6. Morphology.

Section 1.2 has described some of the studies which recorded morphological variation within the *S. damnosum* complex before its sibling status was known, from both East and West Africa, and in adult males, females and larvae. This section will review morphological studies which explicitly aim to distinguish adults of the West African *S. damnosum* complex, rather than those which recorded morphological variation before the recognition of the six main cytospecies (Vajime and Dunbar 1975).

Morphological methods remain the most practicable of all adult identification methods, because they are easy to use and are portable. However, morphological variation includes confounding factors such as environmentally mediated seasonal and/or geographic variation which can give misleading results unless care is taken to sample widely enough. Also, by definition, morphological differentiation between sibling species is not great, so that finding species diagnostic morphological characters is difficult.

a). Soponis and Peterson (1976, also Anon 1976) examined 55 morphological characters on 97 female flies collected from nine sites in Togo. They used univariate statistical methods, examining the sample distributions expressed as histograms for bi- or multimodality, with the flies divided into the three colour categories used by Lewis and Duke (1966). They found that nine characters were unimodal, six characters were bimodal, and 16 characters were multimodal. They found that fore and mid basitarsus length, fore femoral length, wing length, length of the fourth maxillary palp segment and the number of macrotrichia on the radial vein of the wing identified what they believed, based on correlated larval cytotaxonomic identification, to be *S. soubrense* and *S. sirbanum*. Peterson and Dang (1981) subsequently claimed these characters were not practicable for species identification.

b). Quillévére *et al.* (1977) produced a key for the identification of females of the *S. damnosum* complex based on the length of the antenna, the relative compaction of antennal segments 4-8, and the number of maxillary teeth. This key was for the six main species of the *S. damnosum* complex, *S. damnosum s.s.*, *S. sirbanum*, *S. soubrense*, *S. sanctipauli*, *S. squamosum*, and *S. yahense*, although the latter pair could not be distinguished. They did not present formal tables of error rate, so it is not possible to evaluate the power of their key when applied to their own data. Quillévére and Sechan (1978) examined 2468 wings from the same six West African species from five countries, and considered that the number of hairs on the radial vein of the wing could be used to distinguish *S. squamosum* from *S. yahense*, with a certain amount of overlap. This character was included as an extra section in their key.

Garms (1978) evaluated the morphological characters used in this study and confirmed that the length and shape of the antenna was taxonomically useful, but that the number of maxillary teeth and the number of hairs on the radial vein were not. Townson and Meredith (1979) also examined these characters and found that the two doubtful characters were not taxonomically useful because they were significantly correlated with overall size (*contra* Quillévére *et al.* 1977, Quillévére and Sechan 1978), which shows extensive and overlapping variation.

c). Garms (1978) examined seven morphological characters, wing tuft colour, length, shape and colour of the antennae, the number of maxillary teeth, the number of hairs on the radial vein of the wing and the length of the thorax in adult females of six species of the West African *S. damnosum* complex.

He found that the ratio of length of thorax to the length of antenna was a useful taxonomic character as well as wing tuft colour. These characters have been used extensively in subsequent work on the epidemiological significance of different members of the *S. damnosum* complex (e.g. Garms *et al.* 1982, Garms 1983, Cheke and Garms 1983, Garms and Cheke 1985, Cheke and Garms 1986, Cheke *et al.* 1987, Cheke and Denke 1988). The general findings of these papers has been that in the absence of *S. squamosum*, then the species pair *S. sanctipauli/S. soubrense* and *S. sirbanum/S. damnosum s.s.* could be distinguished using the thorax/antennal ratio and wing tuft colour, but *S. squamosum/S. yahense* overlaps with both species pairs when either is present. Garms and Zillman (1984) found a new morphological character (the colour of the setae on the ninth abdominal tergite) which was over 99% diagnostic for *S. yahense* when compared with *S.*

sanctipauli using gel electrophoresis. This character was also used by Thomson *et al.* (1987) who found 91.3% of *S. yahense* had dark abdominal setae.

To conclude, four characters emerged as being taxonomically useful, wing tuft colour, thorax length, antennal length (and shape and colour), and abdominal setal colour. However, the morphometric methods used were of the 'index' kind, rather than using multivariate statistical methods to combine these characters in an optimal way. If multivariate methods had been used, the rate of correct identification would undoubtedly have improved.

d). Dang and Peterson (1980) produced a pictorial key to six species of the West African *S. damnosum* complex, *S. damnosum* s.s., *S. sirbanum*, *S. sanctipauli*, *S. soubrense*, *S. squamosum*, *S. yahense* from an unspecified number of countries. They used 12 characters to identify the adult females and eight to identify adult males.

The characters for identifying adult females were wing tuft colouration, length, shape and colour of the antennae, colour of the scales on the hind leg, colour of the setae on the hind trochanter, colour of the scales of the scutum and the scutellum, colour of the setae of the abdomen, colour of the setae on the vertex, colour of the setae of the postcranium and the colour of the setae on the fore coxa.

The characters for identifying the males were wing tuft colouration, colour of the haltere, colour of the scales of the lateral margin of the scutum, the colour of the setae on the clypeus, the colour of the setae on the postcranium, the extent of the dark spot on the seventh abdominal tergite, and the scutal pattern.

No formal estimate of error rate was presented so it is not possible to evaluate how successful their key was in identifying their own data. Peterson and Dang (1981) extended this work, and showed the distribution of characters in the same six species. They described 18 characters which they considered to be taxonomically important for identifying females, and 13 characters to identify males. However, no estimate of error rate using these characters sets was given making it impossible to evaluate objectively their character sets.

Of the characters described by Dang and Peterson (1980), the male scutal patterns have been used (Meredith *et al.* 1983, Cheke *et al.* 1987), as has the colour of the postcranial hairs (e.g. Walsh *et al.* 1981) and the colour of the scutellar hairs (e.g. Garms 1983).

e). Meredith *et al.* (1983) examined variation in male scutal pattern in the *S. sanctipauli* subcomplex and found considerable variation in this character. Males from Togo and Benin (*S. soubrense* 'Beffa') were predominantly type four, while types one and two (see their figure five) were dominant in the west. They also examined the wing tuft colour character in both sexes, using the categories of Kurtak *et al.* (1981) and found that all five categories were found within *S. soubrense* 'Beffa'.

f). Cheke *et al.* (1987) examined males of *S. sirbanum* reared from pupae collected at 14 sites in four West African countries, evaluating the male scutal pattern described by Dang and Peterson (1980) as being taxonomically important in distinguishing *S. sirbanum* males from *S. damnosum* s.s. males. They considered two hypotheses to explain variation in the scutal patterns that they found between more northerly samples and more southerly samples, either that *S. sirbanum* is

polymorphic or that the that the variation represents two different cytospecies. In support of the second possibility they cite Philips *et al.* (1985) who found significant intraspecific variation in cuticular hydrocarbons between northern and southern *S. sirbanum*. They conclude that the cytotaxonomic status of *S. sirbanum* needs to be further investigated.

g). Recently, Beech-Garwood *et al.* (in the press) have found that the presence of golden hairs on the mesonotum of female *S. damnosum* s.l. in Sierra Leone is a good indicator of the presence of *S. squamosum* in areas where *S. squamosum* and *S. soubrense* or *S. soubrense* 'B' may be sympatric. The character is not diagnostic however, since a minority of *S. soubrense* and *S. soubrense* 'B' may also have these hairs. The observation has been confirmed by enzyme electrophoresis by Davies *et al.* (1988). They conclude that the character is useful at the population level, but needs to be supported by other evidence for single fly identification.

1.4 MORPHOLOGICAL DATA AND MULTIVARIATE STATISTICS

1.4.1 BASIC NOTATION

Multivariate problems are defined by Gnanadesikan (1977) as those concerned with the analysis of n points in p -space, i.e. each of the n objects (in this project, the number of flies) has associated with it a p -dimensional vector of responses (in this analysis, the 28 characters measured or scored on each fly).

The basic difference between this approach and the univariate approach is that the variation in the p -dimensional vector of characters is treated simultaneously, and any information contained in the association (correlation) between characters exploited.

Some of the basic notation and concepts which are needed for an understanding of multivariate statistics can be found in standard texts such as Seber (1984), Mardia, Kent and Bibby (1979) and Gnanadesikan (1977).

If X , is the matrix of n (number of observations) rows by p (number of characters) columns, then the mean vector can be calculated as,

$$\bar{x} = 1/n \sum x_i \dots \dots \dots 1$$

where x_i is the p -dimensional vector of observations on the i -th fly, summation is over $i=1, \dots, n$.

The $n \times p$ matrix of centred observations, \tilde{X} can be calculated,

$$[x_1 - \bar{x}, \dots, x_n - \bar{x}]$$

and from this the $p \times p$ matrix of sums of squares and cross products (SSQPR) is given by,

$$Q = \tilde{X}' \tilde{X} \dots \dots \dots 2$$

The sample dispersion (variance/covariance) matrix is calculated as,

$$S = Q / (n-1) \dots \dots \dots 3$$

The $p \times p$ matrix of Pearson product moment correlation coefficients is derived from the dispersion matrix by first calculating D , the diagonal matrix of variances (i.e. the principal diagonal of S). Then,

$$R = D^{-0.5} S D^{-0.5} \dots\dots\dots 4$$

A basic statistic in multivariate analysis is Mahalanobis' distance (Mahalanobis 1936), which can be calculated as the distance between individuals in a sample, or as the distance between the individuals and the sample mean vector, or as the distance between mean vectors, e.g.,

$$D^2 = (x_i - \bar{x})' S^{-1} (x_i - \bar{x}) \dots\dots\dots 5$$

which is the distance from the i -th individual to the mean vector. The importance of this distance lies in the use of S , the dispersion matrix to weight the distance, thus using the information contained in the variance of the characters and the correlation between them.

1.4.2 ASSUMPTIONS OF MULTIVARIATE ANALYSIS

The use of multivariate statistical methods includes accepting some assumptions about the data set which might not be met. It is usually assumed that the data are sampled from a multivariate normal population with mean \bar{x} and dispersion matrix Σ (Seber 1984). This assumption is rarely exactly met with real data, but if the assumption is made, tested and approximately met then statistical hypotheses about the mean vector and covariance matrix can be performed. If the assumption is made and tested, then more informal data analytic multivariate methods can be used (Gnanadesikan 1977) whether the assumption is met or not. In particular, graphical and dimension reduction techniques can be applied to the data, methods which can also be used on data on which no assumptions are made.

In practise, real morphological data tend to be approximately normally (symmetrically) distributed (Campbell 1978), so the assumption of multivariate normality is often reasonable, although careful checks must be made on the characters, either jointly or separately to ensure that they do conform approximately to normality. Of course, characters which violate the assumption can be used, but not in a formal statistical hypothesis testing regime. The repertoire of nonparametric multivariate statistical methods is limited, so the use of the multivariate normal model is made partly because of the tendency for real data to approximate to it and partly because this is the best developed aspect of multivariate analysis.

In the structured data situation, it is usually assumed that the dispersion matrices of each group (samples, or species) are the same. This is assumed even though differences in covariance structure are one way in which organisms can differ in shape (Reyment 1962, Gould 1984). However organisms can differ in shape but share the same covariance structure (Campbell 1978), so unequal dispersion is not a necessary condition for shape differences to occur. The assumption can be tested (Chapter seven), although these tests are generally inadequate (Seber 1984).

1.4.3 AIMS AND METHODS OF MULTIVARIATE ANALYSIS

The analysis of a multivariate data set can be very complex (Gower and Digby 1981), and many methods have been developed to help in the interpretation of multivariate data (see for example, Gnanadesikan 1977, Gordon 1981, Seber 1984).

The particular multivariate statistical method(s) to be applied to a data set clearly depends on the original objectives of the specific project (Atchley and Bryant 1975). However, the strategy to

be followed can be classified into broad regions which will to some extent dictate the methods of analysis which can be used.

First, the aim of a study might be to test a very specific hypothesis or set of hypotheses (such as testing the null hypothesis of no genetic component to size and shape in the rat, Atchley *et al.* 1982) in which case the appropriate methods include multivariate analysis of variance (MANOVA) and canonical correlation analysis. Or the aim might be to explore the data informally using data-analytic methods, perhaps revealing previously hidden patterns within the data, in which case principal components analysis (PCA), cluster analysis (Gordon 1981), non-metric multidimensional scaling (Kruskal 1977) and graphical methods (Gower and Digby 1981) are more appropriate.

Secondly, the data might be structured and the aim of the study is to explore any differences or similarities within and between data defined by these structures. This could take the form of formal hypothesis testing, using canonical variates analysis (CVA) or MANOVA, or it could be informal exploration, in which case PCA, cluster analysis etc. could be used. Or the data might be unstructured, in which case the aim of the study might be to explain and summarise the observed variation within the data, again either in a formal hypothesis testing regime (using for example canonical correlation analysis) or not (using PCA, cluster analysis, nonmetric MDS etc.).

Clearly these areas are very broad and overlapping, and a particular study is unlikely to adhere strictly to one or other of these strategies, for example, a study which aims to test a specific hypothesis will need to use informal data-analytic methods both to gain a better understanding of the data and to ensure that the assumptions of the model are not violated.

1.4.4 MULTIVARIATE ANALYSIS APPLIED TO MORPHOLOGICAL DATA

Multivariate statistics and biology have developed together (e.g. Weldon 1893, Pearson 1926, Fisher 1936, Rao 1948, Mahalanobis *et al.* 1949, Blackith and Reyment 1971, Sneath and Sokal 1973) because biological data are as a rule inherently multivariate. Multivariate methods do not have to be applied solely to morphological data (e.g. Nei 1987) but in practise this is probably the most commonly analysed form of biological data, at least using methods exploiting character correlations. If more general data are used (e.g. morphology, allozyme frequencies, immunological distance etc.) to infer evolutionary relationships between taxa then the discipline is generally known as numerical taxonomy (Sneath and Sokal 1973). Many of the methods used in numerical taxonomy are cluster analysis techniques applied to proximity matrices calculated between taxa, often not incorporating information about the correlation between the characters. The more restricted discipline, multivariate morphometrics (Blackith and Reyment 1971) in general uses statistics exploiting character correlations, but is applied only to quantifications of morphological characters. Multivariate morphometrics is not necessarily interested in evolutionary relationships in the sense of taxonomies. In practise, the distinction between the two disciplines is not great, (Blackith and Reyment 1971 suggest the term 'quantitative taxonomy' as a more general term, but much multivariate morphometrics is not taxonomic in the strict sense) and studies using both approaches are common (e.g. Lindensfelser 1984).

Insects have often been used in morphometric studies (Daley 1985), their abundance, evolutionary and ecological importance, and the large number of characters which can be measured make them ideal subjects.

Morphometric studies of the Diptera have been common, the following are examples from the literature on morphometrics of the Diptera, organised by family.

In the Culicidae, Rohlf (1963) examined the congruence between larval and adult *Aedes* classifications, using 48 species, 71 larval and 72 adult coded characters. He applied hierarchical cluster analyses to the proximity matrices calculated using the correlation similarity coefficient and a distance metric and compared the resultant dendrograms. He concluded that while there was congruence between the classifications using the two life stages, it was not good enough, and therefore a realistic taxonomy should take account of all life stages simultaneously. Ribiero (1980) examined 34 characters in the *Anopheles gambiae* complex using cluster analysis and ordination methods. Significant morphometric differentiation was found. Dahl *et al.* (1984) used pattern recognition applied to claw shape in mosquitoes, five species of *Aedes*, and one species each of *Anopheles* and *Culex*. Rohlf and Archie (1984) used fourier methods to characterise wing shape in 127 species of North American mosquitoes, which they considered a useful method for quantifying shape variation.

In the Chironomidae, Atchley and Martin (1971) examined sexual dimorphism in 17 larval head capsule characters from five species of *Chironomus*. They used discriminant analysis and canonical variates analysis to compare and contrast patterns of sexual dimorphism within the five species. They found a correlation between the degree of sexual dimorphism in a species and the amount of chromosomal polymorphism (using polytene chromosome analysis), and also that the nature of the dimorphism differed between species, as revealed by CVA ordination (for a similar but more sophisticated analysis of sexual

dimorphism applied to primates see Oxnard 1984). Atchley (1971a) examined sexual dimorphism in five species of *Chironomus* using factor analysis. He found that in most species the dimorphism was along the size axis but in one species dimorphism was along a shape factor, which he explained by hypothesising ecological differences. Titmus and Badcock (1981) examined parasitic feminisation in a Chironomid, *Einfeldia dissidens* resulting from mermithid infection, using CVA, and found two axes of variation, one corresponding to sexual dimorphism and one corresponding to parasitisation. Thus parasitised males and females were closer to each other than were unparasitised and parasitised females, although all these were much closer to each other than any were to unparasitised males.

In the Ceratopogonidae, Atchley (1971b) examined geographic variation in 14 characters in the pupae of three *Culicoides* species, and found that the relative amounts and nature of the variation differed between the three species. Atchley (1971c) extended this work on the three same siblings of *Culicoides* using factor analysis and multiple regression. He found that the proportion of variation which could be accounted for by regression of climatic variables onto morphological characters varied according to species. He explained these differences in terms of Levins' (1965) adaptive models with some species responding to selection (poorly buffered) more than others (well buffered). Atchley (1973) used CVA and stepwise discriminant analysis to separate three species of the *Culicoides* (*Selfia*) group. Starting with 43 characters measured on 298 pupal and adult flies, he derived a subset of 7 characters for overall discrimination, and subsets of 8, 9 and 4 characters for the species-pair analyses. He found significant morphometric differentiation, but considered that

the allocation rate using just adult characters was not good enough, and pupal characters were also needed, limiting the practical use of the method. Atchley (1974) examined 13 characters from 113 adult females of two species of Ceratopogonidae, *Leptoconops torrens* and *L. carteri* using CVA and nonmetric multidimensional scaling. Reducing the initial set of characters from 13 to 6 using stepwise discriminant analysis, he found complete separation between the two species. When further specimens were allocated using this 6 character set, 95% were allocated correctly. Hensleigh and Atchley (1977) examined variation in *Culicoides variipennis* (vector of blue-tongue virus in North America) in laboratory controlled conditions. Their aim was to investigate intraspecific variation which had resulted in the naming of 5 subspecies, using CVA, stepwise discriminant analysis, factor analysis, analysis of variance and multiple regression. By artificially rearing flies at different temperatures they were able to show that most of the variation found in natural populations was due to temperature variation, bringing into question the naming of subspecies. Lane (1981) examined wing spot patterns in the *Culicoides pulicaris* group, using two methods of coding the characters, one without taking account of morphogenetic parameters, the other taking these into account. Principal co-ordinates analysis and principal components analysis were used as ordination methods applied to the two methods. The method which took account of morphogenesis was considered superior in explaining observed variation within the group.

In the Phlebotominae, Lane and Ready (1985) examined six characters in *Lutzomyia wellcomei* and *Lu. complexus* and found significant morphometric differentiation, although there was considerable overlap between the two species.

In the Muscidae and Drosophilidae, Rohlf and Sokal (1972) examined 14 morphological characters in *Musca domestica* and *Drosophila melanogaster* using factor analysis of the correlation matrices. The five factors they extracted from the two species were considered by them to be homologous. Bryant and Turner (1978) examined variation in *Musca domestica* and *M. autumnalis* using the same characters as Rohlf and Sokal (1972). They used principal components analysis and a factor congruence method involving least squares comparison of principal components to examine geographic variation, and found that the patterns of variation were similar in the two species, but that the genetic component of variation in *M. autumnalis* was less than in the house-fly, a finding which they attributed to the recent introduction of the face-fly, resulting in a population bottleneck.

Brown (1979, also Brown and Shipp 1977, 1978) examined wing variation in the Calliphoridae and the Sarcophagidae using CVA and cluster analysis to compare and contrast the taxonomies derived using traditional taxonomic methods with those derived using numerical methods.

Apart from the Diptera, insect morphometrics has included a wide range of families, most notably in the Orthoptera (e.g. Roy and Mukherjee 1964, Blackith and Blackith 1969, Atchley and Hensleigh 1974, Campbell and Dearn 1980), in the Hemiptera and Homoptera (e.g. Sokal and Thomas 1967, Jeffers 1967, Davies and Boryatinski 1979, Bird *et al.* 1981, Simon 1983), in the Coleoptera (Lubischev 1962) and in the Hymenoptera (e.g. DuPraw 1965, Plowright and Stephen 1973), although this is by no means a comprehensive list (see Daley 1985).

Other invertebrates which have been analysed using multivariate statistical methods include bivalves (Ferson *et al.* 1985, Davis 1983),

Foraminifers (Reyment 1982), horseshoe crabs (Riska 1981), Sea Urchins (Lessios 1981), land snails (Gould *et al.* 1975, Gould 1984) and prawns (Lindenfelser 1984).

Vertebrates have been extensively examined using multivariate morphometric methods, including Amphibia, (Reyment 1961), Reptilia (Jolicoeur and Mosimann 1960, Thorpe 1980), Birds (e.g. Schnell 1970, Johnston and Selander 1971, Rising 1970). Within the mammals, bats (e.g. Baker *et al.* 1972, Campbell and Kitchener 1980), rodents (Corbet *et al.* 1970, Thorpe and Leamy 1983, Atchley *et al.* 1982), carnivores (Jolicoeur 1959) and primates (e.g. Mahalanobis *et al.* 1949, Ashton *et al.* 1965, Van Vark and Howells 1984).

1.4.5 COMPUTER PROGRAMS FOR MULTIVARIATE ANALYSIS

While it is possible to calculate some multivariate statistics without a computer, it is impossible to use the full range of statistical methods without the help of a powerful computer and well written software. Fortunately, high speed computers are widely available, and statistical packages have been written to run on these which provide most of the statistical procedures needed in a typical project.

In this project, the following statistical packages were used,

1. SAS (Statistical Analysis System, SAS Institute 1984, 1986, release 5.16) is a comprehensive system for data analysis, offering a very wide range of data management facilities, univariate and multivariate statistical procedures. Graphical procedures are provided by SAS/GRAPH (SAS Institute 1985, version 5), and matrix algebra is provided by SAS PROC MATRIX, providing a facility for developing new procedures or customising other procedures. A powerful macro

facility is provided, and the system can be run interactively using the display manager system.

2. SPSSX (Statistical Package for the Social Sciences, SPSS inc. 1985, ver. 2.2) allows the analysis of data using a wide range of univariate and multivariate statistical methods. It is generally not as flexible as SAS, but is simple to use.

3. GENSTAT (General Statistical Package, Lawes Agricultural Trust 1984, release 4.04B) provides a wide range of univariate and multivariate statistical methods, and is particularly good for the analysis of designed experiments. A macro library is provided, and the ability to write macros using matrix algebraic expressions makes it flexible. However, it is difficult to use and is poorly documented.

4. CLUSTAN (Cluster Analysis Package, Wishart 1978, release 2.1) is a specialised package offering a very wide range of cluster analysis methods, and some graphical procedures. The package is widely used across many disciplines.

5. NTSYS (Numerical Taxonomic System of Multivariate Statistical Programs, Rohlf 1985) is a specialised package allowing cluster analysis, and ordination of numerical taxonomic data. Limited matrix algebraic manipulation is allowed.

1.5 OBJECTIVES OF THE PROJECT

It is clear from the review of the literature in 1.3 that there is no simple set of morphological characters that can be used in traditional taxonomic keys to separate females of all the sibling species of the *S. damnosum* complex in West Africa. Morphometric methods based on multivariate statistical analyses have been used successfully in other groups of insects, so the basic techniques are widely available and well understood.

Therefore, the main objective of this project is to use multivariate statistical techniques to find combinations of morphological characters which can best identify adult females of the *S. damnosum* complex in West Africa.

The characters measured or scored on each adult female fly are described in Chapter four, and statistical methods used to screen the basic data set are described in Chapter five.

Chapter six introduces some of the multivariate statistical methods used for description of morphological variation, and applies these methods to intraspecific variation within cytospecies of the *S. damnosum* complex.

Chapter seven presents the statistics necessary for regional allocation of unknown adult female flies from two regions, Togo and Benin, and the area west of the Volta Lake, Ghana.

Chapter eight describes the statistics necessary for the allocation of unknown females without prior knowledge of the geographic origin of the fly, while the final chapter draws general conclusions about the method of identification, gives worked examples of the mathematics involved and suggests a protocol for the field identification of adult females based on the statistics presented in the

previous two chapters. Details of the data set are presented as an appendix (Appendix one).

Prior to the multivariate morphometric analysis of adult female *S. damnosum s.l.*, it was necessary to obtain as much correlated larval cytotaxonomic identifications as possible. As a result of this work, a new cytotype within *S. sanctipauli* was found from Togo, which is presented as Chapter two. Chapter Three applies multivariate statistical methods to available data within the *S. sanctipauli* subcomplex to examine between and within species variation, the first such analysis of polytene chromosome variation.

CHAPTER TWO: THE CYTOTAXONOMY OF SIMULIUM SANCTIPAULI DJODJI FORM

2.1 INTRODUCTION

The importance of describing genetically distinct forms or geographic races within previously recognised cytospecies of the *S. damnosum* complex comes from the possible correlation of the different cytoforms with factors of significance in disease transmission, such as anthropophily (Cheke and Denke 1988), together with the use of new forms in tracing migration patterns or the distribution of insecticide resistance (Post and Kurtak 1987).

The purpose of this chapter is to describe a new cytotoxic form within *S. sanctipauli* from Ghana and Togo, which was discovered during routine cytotoxic identifications to provide correlated chromosomal identities for adults reared from pupae. The adults were used in the morphometric analyses described in Chapters six, seven, and eight.

2.2 MATERIALS AND METHODS

Breeding sites where the Djodji form of *S. sanctipauli* was collected are listed in Table 2.1. Larvae were fixed in 3:1 ethanol:acetic acid and stored in a refrigerator. For preparation of polytene chromosomes, the larvae were split open ventrally and hydrolysed in 5M Hydrochloric acid for one hour. The ^{salivary}glands were then separated from the larval body and stained in a drop of lacto-propionic orcein (Macgregor and Varley 1983) and mounted in 60% acetic acid. Photographs were taken of each chromosome arm, and the prepa-

ration made permanent by prising off the cover slip (after cooling in liquid nitrogen), immersing the slide in absolute ethanol for one minute then adding a drop of euparal onto the preparation and lowering a new cover slip onto the slide. The slide was then dried on a warm plate for some months, and stored. The larval body was washed in distilled water and put in a glass vial with Feulgen (Macgregor and Varley 1983) until the body had stained. The larval sex was determined using the shape of the developing gonads (Puri 1925). Inversions were scored by comparison with the standard maps of Post (1986), with *S. squamosum* as the reference sequence.

2.3 CHROMOSOMAL CHARACTERISTICS AND CYTOTAXONOMIC KEY

All fixed and polymorphic inversions within Djodji form are indicated on the idiogram (Figure 2.1), and frequencies of polymorphic inversions are listed as Table 2.3. The new form is fixed for the inversions 1L-P&Q, 2L-4&6&A and 3L-2, but there are no new fixed inversions unique to Djodji form, and only one new rare polymorphic inversion (1S-P, see Figure 2.2). The presence of inversion 2L-A places Djodji form in *S. sanctipauli* (Post 1986). However, 1S-21 (Figure 2.3) is strongly linked to the Y-chromosome in Djodji form (Table 2.2), and this unique feature is the most important cytotoxic criterion for both description and routine identification. Since 1S-21 is Y-linked in Djodji form there is no single inversion which is diagnostic of all individuals. However, samples in which there is strong Y-linkage of the inversion can be unequivocally identified as *S. sanctipauli* 'Djodji', and mixed samples (should they exist) of the form with typical *S. sanctipauli*, will be recognised as such using population genetic analysis.

The following cytotaxonomic key can be used for the identification of *S. soubrense* 'Beffa', *S. sanctipauli* and the Djodji form of *S. sanctipauli*. The key should not be used west of Côte d'Ivoire, where typical *S. soubrense* and *S. soubrense* 'B' might also be encountered.

- 1) Larva homozygous for inversions 1L-P&Q, 2L-4&6 and 3L-2
.....*S. sanctipauli* subcomplex 2)

These inversions absent from larva

.....Other species of

S. damnosum complex

- 2) Larva homozygous for inversion 2L-A
.....*S. sanctipauli* 3)

Inversion 2L-A absent from larva

.....*S. soubrense* 4)

- 3) Inversion 1S-21 Y-linked in population
.....*S. sanctipauli* 'Djodji'

Inversion 1S-21 not Y-linked in population

.....*S. sanctipauli* typical

- 4) Inversion 2S-6b present in larva
.....*S. soubrense* 'Beffa'

Inversion 2S-6b absent from larva

.....*S. soubrense* typical

2.4 DISCUSSION

The new cytotype seems to be largely limited to the Asukawkaw and Dayi river systems in the mountainous forest on the Ghana/Togo border (see Table 2.1). Within Togo and Benin, to the north and east, *S. soubrense* 'Beffa' appears to be the sole representative of the *S. sanctipauli* subcomplex except for a few samples of the Djodji form identified further north in the savanna from the rivers Kpaza and Niankpe in October 1987. To the west of the Volta lake *S. sanctipauli* typical form and *S. soubrense* are found (Meredith *et al.* 1983, Post 1986, Fiasorgbor, Weber, Post, Surtees unpublished data, Chapter three).

In view of the absence of any sympatric samples or unique fixed inversions, there is no evidence for Djodji form being a species distinct from *S. sanctipauli* elsewhere. However, the sex-linkage of 1S-21 shows that Djodji populations are by definition genetically differentiated from other *S. sanctipauli* populations. There is also evidence for multivariate morphometric differentiation between typical *S. sanctipauli* and *S. sanctipauli* 'Djodji' (Chapter six). Therefore it seems that Djodji form should be considered to be a geographic race of *S. sanctipauli*.

The taxonomic significance of sex-linked inversions in the Simuliidae has been discussed by Rothfels (1979), Rothfels and Nambiar (1981) and Post (1982). Most blackfly species do not have distinguishable sex chromosomes, but sex chromosome differentiation can occur, often by linkage of inversions to the primary sex-determining region. Often species differ only in their sex chromosomes, which has led to the hypothesis that sex chromosome evolution may play a functional role in speciation within blackflies (Rothfels 1979).

Within the *S. damnosum* complex sex-linked inversions have been considered important in the cytotoxic description of several forms and species, such as *S. soubrense* 'Beffa' (Meredith *et al.* 1983), *S. yahense* (Vajime and Dunbar 1975), and Turiani form (Dunbar and Vajime 1981).

The importance of Djodji form lies in its possible importance in onchocerciasis transmission, which is discussed by Garms and Cheke (1985) and Cheke and Denke (1988). Cheke and Denke (1988) show that *S. sanctipauli* 'Djodji' is potentially a better vector than *S. squamosum* in Togo, and better than *S. sanctipauli* from Côte d'Ivoire, where *S. sanctipauli* is believed to be more zoophilic.

List of larval samples from which the Djodji form of *S. sanctipauli* has been identified.

| Sample | River | Coordinates (N/E) | Date | Collectors ¹ | Cytospecies ² composition | | | | |
|--------|-----------|----------------------|----------|-------------------------|---|-----|-----|-----|-----|
| | | | | | squ | yah | san | dam | sir |
| 1 | Dayi | 7°09' 0°29' | 13.01.87 | AKA | 22 | 5 | 3 | 1 | 0 |
| 2 | Dayi | 7°07' 0°27' | 22.01.87 | RAC EAW YY | 20 | 74 | 8 | 1 | 0 |
| 3 | Dayi | 7°06' 0°26' | 17.03.86 | AKA | 31 | 12 | 4 | 0 | 0 |
| 4 | Dayi | 7°06' 0°26' | 29.04.86 | SSH | 20 | 5 | 4 | 0 | 0 |
| 5 | Dayi | 6°57' 0°21' | 12.02.86 | AKA | 17 | 15 | 21 | 55 | 2 |
| 6 | Dayi | 6°57' 0°21' | 21.03.86 | AKA | 9 | 0 | 8 | 47 | 0 |
| 7 | Dayi | 6°53' 0°21' | 06.05.86 | SSH | 2 | 0 | 4 | 8 | 3 |
| 8 | Dayi | 6°53' 0°21' | 16.05.86 | SSH | 5 | 2 | 17 | 16 | 8 |
| 9 | Dayi | 6°53' 0°21' | 13.01.87 | RAC EAW YY | 2 | 0 | 4 | 8 | 3 |
| 10 | Dayi | 6°52' 0°19' | 29.05.86 | AKA | 2 | 0 | 8 | 19 | 2 |
| 11 | Dayi | 6°52' 0°19' | 22.01.87 | RAC EAW YY | 2 | 0 | 1 | 7 | 0 |
| 12 | Asukawkaw | 7°54' 0°37' | 05.02.87 | JFW JEEH | 54 | 0 | 7 | 0 | 0 |
| 13 | Asukawkaw | 7°54' 0°37' | 27.05.86 | YY | 41 | 0 | 13 | 0 | 0 |
| 14 | Asukawkaw | 7°54' 0°37' | 23.01.87 | RAC EAW YY | 91 | 0 | 27 | 0 | 0 |
| 15 | Asukawkaw | 7°52' 0°36' | 18.03.86 | MD | 26 | 0 | 4 | 0 | 0 |
| 16 | Asukawkaw | 7°52' 0°29' | 23.01.87 | RAC EAW YY | 33 | 0 | 28 | 0 | 0 |
| 17 | Asukawkaw | 7°41' 0°26' | 28.05.86 | YY | 29 | 0 | 52 | 0 | 0 |
| 18 | Asukawkaw | 7°41' 0°25' | 06.02.86 | | 5 | 0 | 90 | 0 | 0 |
| 19 | Asukawkaw | 7°41' 0°25' | 23.01.87 | RAC EAW YY | 2 | 0 | 74 | 0 | 0 |
| 20 | Menou | 7°37' 0°39' | 28.05.86 | YY | 28 | 0 | 14 | 0 | 0 |
| 21 | Gban-Houa | 7°42' 0°38' | 28.05.86 | YY | 3 | 0 | 25 | 0 | 0 |
| 22 | Gban-Houa | 7°42' 0°38' | 23.01.87 | RAC EAW YY | 12 | 0 | 29 | 0 | 0 |
| 23 | Gban-Houa | 7°41' 0°37' | 06.02.86 | YY | 12 | 0 | 18 | 0 | 0 |
| 24 | Gban-Houa | 7°42' 0°36' | 23.01.87 | RAC EAW YY | 11 | 0 | 22 | 1 | 0 |
| 25 | Gban-Houa | 7°42' 0°35' | 15.10.84 | RJP CKP | 4 | 0 | 49 | 0 | 0 |
| 26 | Gban-Houa | 7°42' 0°35' | 15.03.85 | RAC AMD | 17 | 0 | 23 | 0 | 0 |
| 27 | Gban-Houa | 7°42' 0°35' | 21.03.85 | RAC AMD | 11 | 0 | 23 | 0 | 0 |
| 28 | Gban-Houa | 7°42' 0°35' | 26.03.85 | RAC AMD | 8 | 0 | 32 | 0 | 0 |
| 29 | Gban-Houa | 7°42' 0°35' | 29.03.85 | RAC AMD | 10 | 0 | 29 | 0 | 0 |
| 30 | Gban-Houa | 7°42' 0°35' | 15.10.85 | RAC AMD | 13 | 0 | 21 | 0 | 0 |
| 31 | Gban-Houa | 7°42' 0°35' | 15.03.86 | YY | 45 | 0 | 22 | 0 | 0 |
| 32 | Gban-Houa | 7°42' 0°35' | 27.01.87 | RAC HSA | 28 | 0 | 54 | 1 | 0 |
| 33 | Wawa | 7°43' 0°33' | 20.03.86 | YY | 13 | 0 | 12 | 0 | 0 |
| 34 | Wawa | 7°43' 0°33' | 23.01.87 | RAC EAW YY | 13 | 0 | 32 | 0 | 0 |
| 35 | Wawa | 7°41' 0°30' | 27.03.86 | AKA | 17 | 0 | 6 | 0 | 0 |
| 36 | Kpaza | 8°33' 0°41' | 15.10.87 | JFW SS | | | | | |
| 37 | Kpaza | 8°33' 0°41' | 22.10.87 | JFW YY AKO | | | | | |
| 38 | Kpaza | 8°33' 0°37' | 22.10.87 | JFW YY AKO | | | | | |
| 39 | Niankpe | 9°05' 0°42' | 22.10.87 | JFW YY AKO | | | | | |

¹ AKA=A.K.Adzah, AKO=A.K. Opoku, AMD=A.M. Denke, CKP=C.K. Post EAW=E.A. Weber, HSA=H.S.K. Avissey, JEEH=J.E.E. Henerickx, JFW=J.F. Walsh, MA= M. Ampah, MD=M. David, RAC=R.A. Cheke, RJP=R.J. Post, SS=S. Sowah, SSH=OCP subsector Hohoe, YY=Y. Yamagata. Samples 25-30 were determined by D.P. Surtees.

² sq= *S. squamosum*, yah=*S. yahense*, san=*S. sanctipauli* 'Djodji', dam=*S. damnosum* s.s., and sir=*S. sirbanum*. Other cytospecies were not found in these samples, except for 1 *S. soubrense* in sample 37. Samples 36-39 were not random samples, so numbers of flies identified for each species are not given.

Table 2.2

1S-21 karyotype frequencies

| River | Samples ¹ | Number of males | | | Number of females | | |
|-----------|----------------------------|-----------------|-------|-------|-------------------|-------|-------|
| | | st/st | st/21 | 21/21 | st/st | st/21 | 21/21 |
| Dayi | 1,2,4,7,8 9,11 | 2 | 14 | 0 | 22 | 1 | 0 |
| Asukawkaw | 14,16,19 | 0 | 58 | 0 | 31 | 2 | 0 |
| Gban-Houa | 22,24,26,27 28,29,30,32 | 3 | 108 | 0 | 113 | 1 | 0 |
| Wawa | 34 | 1 | 12 | 0 | 8 | 0 | 0 |
| Kpaza | 36,37 | 0 | 10 | 0 | 11 | 0 | 0 |

Table 2.3

Autosomal polymorphic inversion frequencies

| River | Samples ¹ | Polymorphic inversions ² | | | | | |
|-----------|----------------------------|-------------------------------------|------|------|---------|-------|-------|
| | | 1S-A | 2L-7 | 3L-B | 3L-4.17 | 3L-24 | IS-P |
| Dayi | 1,2,4,7,8 9,11 | 0.56 | 1.00 | 0.0 | 1.00 | 0.0 | 0.0 |
| Asukawkaw | 14,16,19 | 0.42 | 1.00 | 0.0 | 0.97 | 0.0 | 0.0 |
| Gban-Houa | 22,24,26,27 28,29,30,32 | 0.45 | 1.00 | 0.0 | 0.996 | 0.002 | 0.002 |
| Wawa | 34 | 0.50 | 1.00 | 0.0 | 1.0 | 0.0 | 0.0 |
| Kpaza | 36,37 | 0.31 | 1.00 | 0.0 | 1.00 | 0.0 | 0.0 |

¹ Samples are as listed in Table 2.1. However, it was not always possible to score all inversions in every specimen, and hence sample sizes may be slightly smaller than those listed in Table 2.1.

² Inversions 2L-7 and 3L-B were noted heterozygously in a very few specimens from other samples which were not scored systematically for autosomal inversions.

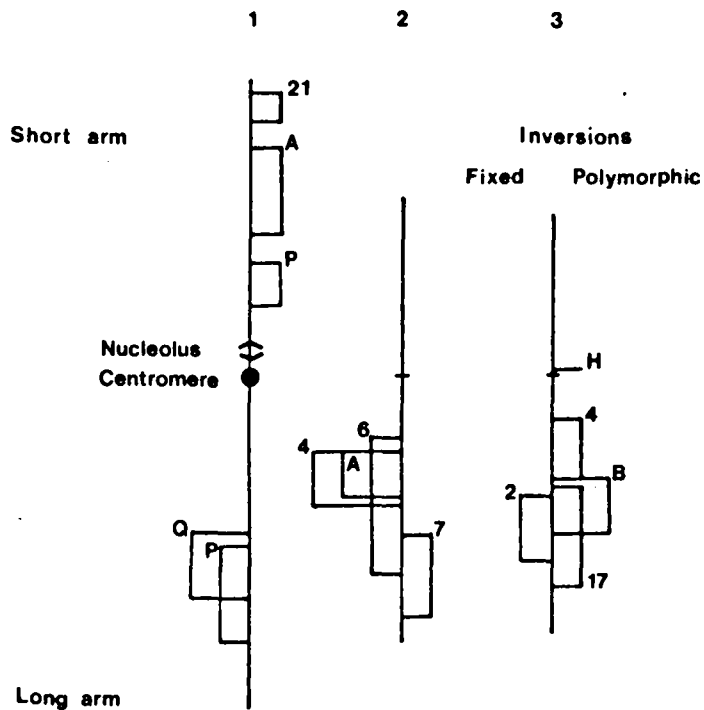


Figure 2.1

Idiogram showing the relative positions of all the inversions currently known from the Djodji form of *S. sanctipauli*. Inversions plotted on the right of each chromosome are intraspecific polymorphisms, while those to the left are fixed inversions relative to the standard sequence in *S. squamosum*. The polymorphic inversion 3L-B is based on the 3L-4.17.2 sequence. Most of these inversions are illustrated in Post (1986), although 1S-21, 1S-A and a new inversion, 1S-P are also shown on Figures 2.2 and 2.3.

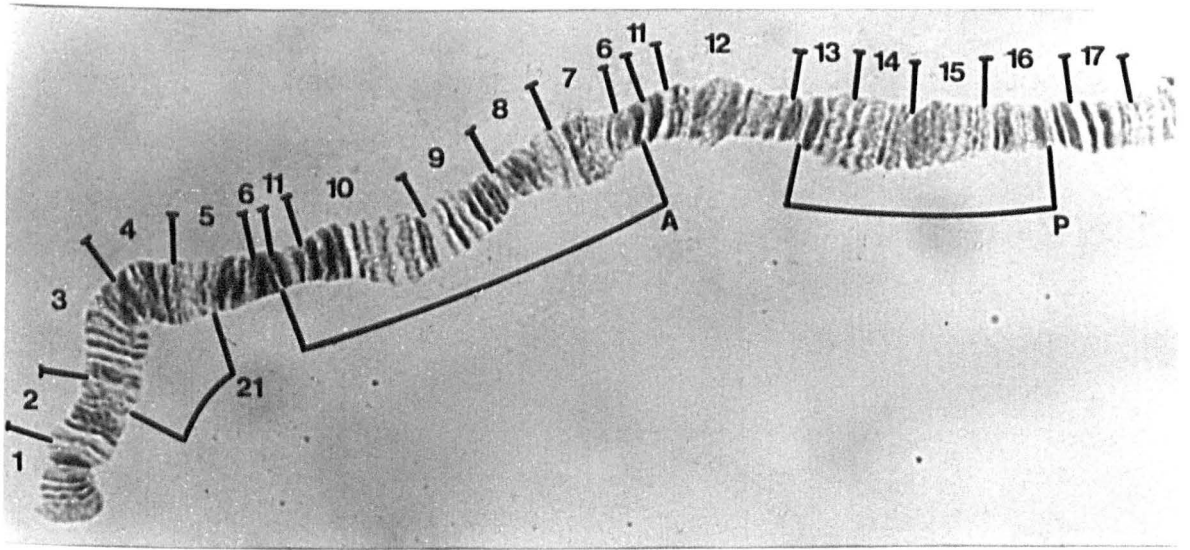


Figure 2.2

S. sanctipauli chromosome arm 1S, from the River Sassandra at Soubre, Cote d'Ivoire, 26.10.84 (see Table 3.1) showing the karyotype 1S-A/A, with the breakpoints of 1S-P and 1S-21 indicated.

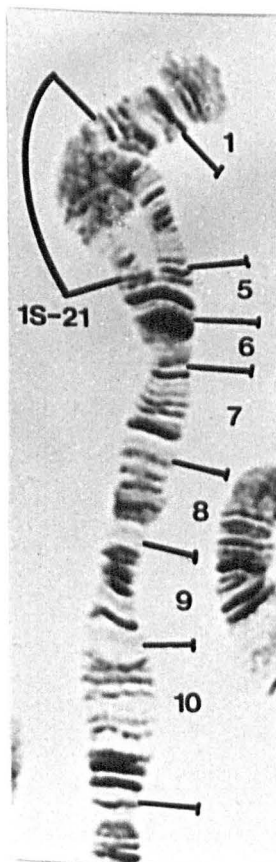


Figure 2.3

S. sanctipauli 'Djodji' chromosome arm 1S, from the river Gban-Houa, Djodji, Togo, showing the 1S-St/21 karyotype.

CHAPTER THREE: MULTIVARIATE ANALYSIS OF POLYTENE CHROMOSOME INVERSION
FREQUENCIES WITHIN THE SIMULIUM SANCTIPAULI SUBCOMPLEX

3.1 INTRODUCTION

The delimitation of taxa within sibling species complexes of medical significance is an important activity before investigation of behavioural, ecological or physiological aspects relevant to disease transmission (WHO Technical Report 597, 1976). Within the *S. damnosum* complex in West Africa this delimitation of taxa was established by Vajime and Dunbar (1975), with later revision of the *S. sanctipauli* subcomplex by Post (1986).

Chapter two has emphasised the essential role that 'classical' cytotoxic methods play in the further development of the understanding of a species complex. The purpose of the present chapter is to use multivariate statistical techniques to analyse, objectively, polytene chromosome variation within and between members of the *S. sanctipauli* subcomplex, which is the best understood part of the *S. damnosum* complex (Post 1986), and to contrast the results thus obtained with those obtained using a 'classical' approach.

The cytotoxicology of the *S. sanctipauli* subcomplex has been split into a number of cytoforms within the three cytospecies, *S. sanctipauli*, *S. soubrense* and *S. soubrense* 'B'. These are:

- S. soubrense* 'Beffa' (Meredith *et al.* 1983)
- S. soubrense* 'Menankaya' (Boakye personal communication)
- S. soubrense* 'Chutes-Milo' (Boakye personal communication)
- S. soubrense* 'Konkoure' (Quillévééré *et al.* 1982)

S. sanctipauli 'Djodji' (Surtees *et al.* 1988, Chapter two)

The usual approach to cytotaxonomy uses a sliding scale of chromosomal evidence for defining taxa (Rothfels 1956, Bedo 1977). The strongest evidence for specific status comes from differences between individuals maintained in sympatry, without apparent introgression (Mayr 1970). As an example from the *S. damnosum* complex, *S. squamosum*, and *S. sanctipauli*, in the River Gban-Houa at Djodji, Togo, have the fixed inversions 1L P&Q, 2L4&6 and 3L-2 maintained between them in sympatry. There are many such examples within the *S. damnosum* complex (Vajime and Dunbar 1975, Dunbar and Vajime 1981, Post 1986). The next strongest evidence for specific status comes from sex-linked inversion differences maintained in sympatry. X chromosome linked inversion differences are as diagnostic as fixed differences, while Y chromosome differences render females of the putative species homosequential. The next strongest evidence comes from polymorphic autosomal inversions found at different frequencies in sympatry. Analysis of sample inversion frequencies, testing for departures from Hardy-Weinberg equilibrium and for linkage disequilibrium can reveal the existence of non-introgressing species in sympatry.

In the absence of sympatric samples, then by analogy, the same three degrees of chromosomal evidence have been used to define taxa in allopatry (for example *S. sanctipauli*, and *S. sanctipauli* 'Djodji', Chapter two), although such evidence for specific status is always weaker than the equivalent evidence in sympatry (Mayr 1970).

When the taxonomy of a group reaches the state of knowledge achieved in the *S. sanctipauli* subcomplex, there is often subjectivity in the criteria used for defining a distinctive taxonomic or cytotaxonomic category and in the taxonomic rank to which a new

cytoform should be raised. Clearly it is advantageous for more rigorous objective methods to be used to describe inter- and intra-specific variation.

Two approaches could be taken to the objective analysis of polytene chromosome variation. One method is to use standard genetic distance models to assess the evolutionary relationships between divergent taxa in a formal hypothesis-testing regime (Nei 1987). This approach has often been used for the analysis of isoenzyme variation, and could be used with cytotaxonomic data. However, within the *S. sanctipauli* subcomplex, the nature of the sampling regime and the sample sizes obtained preclude the use of this approach. Therefore it was decided to use the relatively informal methods of multivariate exploratory data analysis (Gnanadesikan 1977, Gordon 1981), using sample inversion frequencies as continuous pseudo-phenetic characters bounded by zero and one, to analyse inter- and intra-specific variation. Because the taxonomy of this subcomplex is the best known within the *S. damnosum* complex, a well established *a priori* taxonomy based on classical polytene chromosome analysis was available for comparison with the results of the multivariate analysis.

The objectives of this chapter can be clearly stated:

1. To analyse polytene chromosome inversion frequencies in the whole *S. sanctipauli* subcomplex;
2. To compare and contrast the *a priori* taxonomy of the *S. sanctipauli* subcomplex with that derived using multivariate statistical methods;
3. To analyse variation within selected *a priori* defined species.

3.2 MATERIALS AND METHODS

Samples of larvae were available from collections made between 1971 and 1986 from sites between Guinea and Nigeria in West Africa. Table 3.1 gives the details of each sample, the river, site and country of collection, the latitude and longitude of the site, the name of the cytotaxonomist who scored the inversion frequencies and the *a priori* taxonomic category of the sample. Table 3.2 shows the sample inversion frequencies for the 47 inversions variant within the *S. sanctipauli* subcomplex, grouped by chromosome. Sample sizes listed are the maximum. It was not always possible to score all inversions for all specimens so sample sizes for each inversion are occasionally smaller than the maximum.

The chromosome preparations of the samples listed in Table 3.1 were made using the methods described in Chapter two. The larvae were identified chromosomally using the chromosome standards of Post (1986). Most of the identifications were performed by Dr R.J. Post, with others performed by D.P. Surtees.

Two sets of multivariate statistical techniques were used in the exploratory data analysis of the sample inversion frequencies listed in Table 3.2, ordination methods and cluster analysis methods (Gordon 1981). These two sets of techniques were used in a complementary way, with the results of one set helping in the interpretation of the results of the other set (Kruskal 1977). Information in a two dimensional ordination is often easier to interpret if the results from a cluster analysis of the same data are used in conjunction, and similarly an ordination can be used to assess a partition of the data resulting from a particular cluster method.

Within these two sets of techniques different methods were used, rather than using only one ordination method and one cluster analysis method to help to avoid artefacts. Cluster analysis methods in particular can sometimes produce interpretable results even in the absence of real structure in a data set, and can impose a structure on a data set which is different from that actually contained in the data (Gordon 1981). Using a plurality of methods, satisfying different criteria helps to avoid this problem.

For general reference to the methods used in this analysis, see Gordon (1981), Seber (1984), Everitt (1978), Gnanadesikan (1977), Sneath and Sokal (1973).

3.2.1 ORDINATION METHODS

The raw data matrix (Table 3.2) is difficult to interpret as it stands, as it is 66 rows (samples) by 47 columns (inversions) in size. Ordination methods attempt to derive a lower dimensional graphical representation of high dimensional data, which retains as much of the information contained in the full data set as possible while producing a great simplification of the data. There are many ordination methods (Gordon 1981, Seber 1984), including principal components analysis and non-metric multidimensional scaling, which were used in this analysis. The former is an R-mode method, i.e the analysis is performed on the relationships between the characters (inversions) while the latter is a Q-mode method, i.e the analysis is performed on the relationship between individuals (samples of larvae), (Sneath and Sokal 1973).

3.2.1.1 Principal Components Analysis

PCA is a very well established technique (Pearson 1901, Hotelling 1933) which involves the extraction of eigenvalues and eigenvectors

from the sample dispersion (variance/covariance) matrix, or from the correlation matrix derived from the dispersion matrix (Seber 1984). The first principal component is the linear combination of the original variables which accounts for the largest proportion of the total variance, so the first principal plane resulting from the scatter of points in the plane of the first two principal components accounts for the largest proportion of variance of all orthogonal axes. There are many uses to which PCA can be put, including interpreting patterns of covariation between characters, identifying redundant dimensions and identifying outliers (Seber 1984), but in this analysis the method was used as a low dimensional representation of the high-dimensional data set. By maximising variance it is assumed that information content is also maximised in the first few dimensions. In theory and practice this is not necessarily so (Cheng 1983), and so the results of a principal components analysis have always to be interpreted with caution.

Principal components can be extracted from either the dispersion matrix or the correlation matrix derived from it (Seber 1984, Gnanadesikan 1977). In this analysis the dispersion matrix was used, because the first two principal components accounted for a larger proportion of total variance than did the first two principal components of the correlation matrix (see section 3.3). The program used was SAS PROC PRINCOMP.

As an aid to understanding the relationship between the original inversion frequencies and the derived variables (principal components) a graphical technique called h-plotting was used (Corsten and Gabriel 1976, Seber 1984). With this method, each coefficient in the first two principal components (Table 3.3) was multiplied by

the square root of the corresponding eigenvalue. The resultant coordinates show the strength of the relationship between the inversion and the principal components, and when superimposed on an ordination can help to explain the patterns uncovered in the data. The length of each vector is proportional to the standard deviation of the inversion and the cosine of the angle between any two vectors approximates to the correlation coefficient between the inversions. These h-plots were superimposed on the ordinations, but were not scaled to conform to the principal axes, and were not located at the origin as this would have obscured other details of the ordination.

3.2.1.2 Non-Metric Multidimensional Scaling

PCA is an R-mode method because the principal components are extracted from a matrix describing relationships between characters and individuals (samples) are then examined in the space defined by the new linear combinations of the original variables. Non-metric MDS is a Q-mode technique, in that a lower dimensional ordination of the data is derived from a matrix describing the relationship between individuals (Gordon 1981). This matrix could be defined by many proximity measures (Wishart 1978), but in this analysis the squared euclidean distance between individual samples was calculated, because it is the proximity measure also used in the cluster analysis methods described in section 3.2.2. Non-metric MDS works by finding a p-dimensional solution to the transformation of the proximity matrix (in this analysis the 66 by 66 matrix of euclidean distances between samples listed in Table 3.1) to a p-dimensional scatter of points in which the rank order of inter-point distances in the derived space matches as closely as possible the rank order of distances in the original proximity matrix (Gordon 1981). P (the dimensionality of

the final solution) is a parameter defined by the user, and the adequacy of a particular p-dimensional solution is assessed using a parameter called stress (Gordon 1981) which is minimised. The algorithm used in this analysis was an iterative least-squares method, within a general multidimensional scaling package ALSCAL (Young *et al.* 1980).

A problem found with non-metric MDS is the choice of p, the dimensionality of solution. The usual method of choosing p is to plot the change in stress against increasing values of p, and to choose the lowest value of p for which the stress is tolerable. In this analysis, however, if a two-dimensional ordination did not have an acceptable stress value, higher dimensional solutions were not attempted, as this would have defeated the principle behind using the method. In practise, none of the ordinations resulting from the application of this method were used in the final interpretation of the data, because the ordinations resulting from PCA were usually clearer, and the principal axes are more easily interpretable in terms of the original inversions via the eigenvectors and the h-plot.

3.2.2 CLUSTER ANALYSIS METHODS

Ordination methods do not impose a structure onto the data matrix, instead they are a parsimonious representation of high dimensional data. Cluster analysis methods, however, impose a structure onto the data matrix (Gordon 1981). This structure can be of four types:

1. Non-overlapping partitions of the data
2. Hierarchically nested partitions of the data
3. Overlapping hierarchical partitions of the data
4. Overlapping non-hierarchical partitions of the data.

Of these four structures, methods resulting in the first three structures were used in this analysis.

3.2.2.1 Non-overlapping Partitioning Methods

In these methods the data set of n objects is divided into g -groups, where g can be set automatically or by the user depending on the algorithm used (Gordon 1981). Initially, for a g -partition of the data, the objects may be assigned randomly to the g -groups or as a result of a previous cluster analysis method. Objects are then iteratively relocated from one group to another, if doing so helps improve the parameter being optimised. The parameters optimised and the relocation procedure used are all algorithm dependent. In this analysis SAS PROC FASTCLUS was used, which iteratively minimises the sum of squared distances from the g -cluster means.

3.2.2.2 Hierarchical Cluster Methods

Hierarchical methods begin with n clusters (in this analysis n is the number of independent samples i.e. 66) which are successively fused until all belong to one cluster. Clusters formed at a lower level are completely incorporated in higher level clusters (Gordon 1987). In this analysis, four hierarchical cluster methods were used: Single linkage (nearest neighbour), Complete linkage (furthest neighbour), Group Average (UPGMA) and Ward's error sums-of-squares method (Ward 1963). These four were chosen as they are the commonest methods in use, and hence are the best understood (Gordon 1981). In addition, the four methods have quite different statistical properties which accords with the general principle of this analysis that greater confidence is obtained from similar results derived from methods satisfying different criteria.

The different methods work by fusing at each stage the two clusters which are most similar, but they differ in the way each method defines cluster similarity. Single linkage calculates the distance

between two clusters as the distance between the nearest neighbours. Complete linkage calculates it as the distance between the remotest members of the two clusters. Group Average is intermediate between these methods in calculating the distance between two clusters as the average of all pairwise distances between members of the two clusters. Ward's method defines the distance between clusters as the increase in within-groups sums-of-squares which would result from the fusion of two clusters. In all cases, the two clusters for which the various definitions of distance is a minimum are fused (Gordon 1987). Hierarchical methods are usually expressed as two-dimensional branching trees called dendrograms. Within this analysis, the cluster analysis package CLUSTAN (Wishart 1978) and SAS PROC CLUSTER were used.

3.2.2.3 Overlapping Methods

The two previous methods share the restriction that an object can belong to only one cluster. Jardine (1971) argues that certain types of natural variation, including intraspecific geographic variation, is often not appropriately expressed either as an hierarchy or as a non-overlapping partition, but instead may take the form of a continuum (e.g. clinal variation, Endler 1977) or as recognised 'types' with intermediates between them.

To analyse within-taxon variation of the *a priori* taxa *S. sanctipauli*, *S. soubrense* 'Menankaya/Konkoure' and *S. soubrense* 'B' a set of overlapping cluster methods was used. These were the B_k methods of Jardine and Sibson (1968). These methods can best be understood in terms of two parameters, h and k . h is the distance between points and k is the amount of overlap allowed by the method, $k-1$ points being allowed to belong to the overlap of two clusters at a particular value of h . The method starts with $h=0$ (i.e. identical

objects), h is then increased to a particular level (chosen in retrospect by the user because of perceived discontinuities in the parameter, or for other data-analytic reasons). At this value of h , all maximally complete subgraphs are drawn in (i.e. subsets of points in which all the points are connected), and any pair of subgraphs which coincide in at least k points are further fused to form a cluster. For example, if $k=2$, and two clusters share only one point in common, then both clusters remain distinct, but the point in common lies on the overlap of the two clusters. These methods become extremely complex to interpret for moderate numbers of individuals and values of k greater than 4, so higher values of k were not attempted. The CLUSTAN procedure KDEND (Wishart 1978) and the NTSYS procedure BKGRAPH (Rohlf 1984) were used for the analysis.

3.2.2.4 Other Methods

One problem associated with lower dimensional ordinations of high dimensional data is that certain of the interpoint distances become distorted (Seber 1984). To assess this distortion a minimum spanning tree (MST) was calculated and superimposed on the ordination. A MST is the tree connecting the n vertices (in this analysis $n=66$, the number of samples in Table 3.1) forming a connected graph containing no loops for which the sum of the edge lengths is a minimum (Gower and Ross 1969).

Hierarchical techniques can be seen as a transformation of the proximity matrix between individuals into a new proximity matrix satisfying the ultrametric inequality (Jardine and Sibson 1971, Gordon 1981). Generally this transformation results in some distortion, the extent of which can invalidate a hierarchical representation of the proximity matrix. In this analysis, four distortion measures were

used to assess the extent of the distortion: the cophenetic correlation coefficient, r_{cop} (Sokal and Rohlf 1962) and three of the Jardine and Sibson (1968) Δ_i distortion measures.

A common problem to clustering methods is the objective estimate of the true number of clusters within a data set. For hierarchical methods this can be thought of as estimating the level at which a line should be drawn across the dendrogram, and for non-overlapping partition methods this amounts to choosing the value of g at which to stop the algorithm. For this analysis the cubic clustering criterion (CCC) within SAS PROC CLUSTER was used. This criterion compares a particular partition of the data with that expected if the data were sampled from a uniform distribution. The value for which this criterion is largest is taken as the 'true' number of clusters. This method was evaluated by Milligan and Cooper (1985) who found that it compared favourably with most other criteria in the literature, and was considered the best of the widely available methods.

Once a particular partition was obtained by a cluster method it was then compared with results obtained from other cluster methods. A combination of visual inspection of the hierarchical dendrograms to identify cohesive, isolated clusters and a cluster intersection method described in Gordon (1981) to find the maximum number of points in common to two methods was used to compare the results of cluster analyses of the data set. This procedure resulted in a consensus partition of the data defining clusters consistently uncovered using the different methods, but from which certain points were excluded because of their inconsistent classification using different clustering methods.

3.2.2.5 Summary of Methods

Ordination methods and cluster analysis methods were used to analyse the total data set presented in Table 3.2. Principal components analysis and non-metric multidimensional scaling were both used, but only the the results of the favoured method (PCA) are presented. The full data set was clustered using four hierarchical cluster methods and one non-overlapping cluster method. The optimal partition of the data as determined by the CCC was derived for each method and these partitions compared using a cluster intersection method and visual inspection, and a consensus partition of the total data set derived.

In addition to this whole data set analysis, three sub-analyses were performed on the three *a priori* groups *S. soubrense* 'B', *S. sanctipauli* and *S. soubrense* 'Menankaya/Konkoure'. A separate analysis was not performed for *S. soubrense* 'Beffa' or *S. soubrense* 'Chutes Milo' because of the small number of samples of each. An overlapping cluster analysis method was used for these intraspecific analyses as this had been shown in previous studies (Jardine 1971) to be more sensitive to intraspecific variation.

3.3 RESULTS

The first principal plane of the dispersion matrix accounted for 65% of total variance and was used in preference to the first principal plane of the correlation matrix which only accounted for 24% of variance. This was so because many inversions were not correlated, but some inversions had a relatively large variance, the effect of which is damped if the correlation matrix is used in a PCA. A two dimensional solution to a non-metric MDS of the squared euclidean distance matrix between samples resulted in a stress of 12.6%, which is only a 'fair to poor' fit (Kruskal 1964). Therefore this ordination was not used, but the scatter of points in the first principal plane of the dispersion matrix used instead (Figure 3.1). This figure is annotated with the sample numbers corresponding with those in Table 3.1. The points are connected by the minimum spanning tree derived from the squared euclidean distance matrix between samples. Also shown is the h-plot of the dispersion matrix. Table 3.3 gives the first two principal components of the dispersion matrix, demonstrating which sets of inversions have the most influence on the first two principal axes.

Figure 3.2 is the dendrogram resulting from application of single linkage cluster analysis to the squared euclidean distance matrix. The numbers at the tips of the dendrogram correspond to the sample numbers in Table 3.1. Figures 3.3 to 3.5 are the dendrograms resulting from application of complete linkage, Group Average, and Ward's method of cluster analysis.

Table 3.4 gives the measures of distortion (cophenetic correlation coefficient and the three Jardine-Sibson distortion measures) resulting from the hierarchical cluster methods. Also shown is the

partition for each hierarchical method and the non-overlapping partition method for which the CCC was a maximum.

Table 3.5 shows the five partitions resulting from the cluster methods at the level suggested by the CCC for each method, while Table 3.6 is the 'consensus' partition obtained by the method described in section 3.2.2.4.

Table 3.7 gives the reduction in the Jardine-Sibson distortion measure $\Delta_{0.5}$ as the B_k methods were applied to the three data sets analysing the pre-defined groups each of which represents a single species defined by classical cytotaxonomy, *S. soubrense* 'Menankaya/Konkoure', *S. soubrense* 'B' and *S. sanctipauli*. K was increased from one (where it is equivalent to single linkage) to four (beyond which the results were extremely complex).

Figure 3.6 is the first principal plane of the dispersion matrix for *S. soubrense* 'B', with the h-plot superimposed and the cluster boundaries defined by applying the $B_{k=2}$ method at the level $h=0.001$. Figure 3.7 is the scatter of points in the first principal plane of the dispersion matrix of the *S. sanctipauli* samples, with the h-plot superimposed and the cluster boundaries defined by applying the $B_{k=2}$ method at $h=0.001$. Figure 3.8 is the first principal plane of the dispersion matrix for *S. soubrense* 'Menankaya/Konkoure' with the h-plot superimposed and the cluster boundaries defined by applying the $B_{k=2}$ method at level $h=0.001$.

3.4 DISCUSSION

3.4.1 SIMULIUM SOUBRENSE 'B'

Simulium soubrense B (samples 32-45) is the most clearly defined and internally most homogeneous taxon, both on the ordination (Figure 3.1) and on the dendrograms resulting from the application of the four hierarchical cluster methods (Figures 3.2-3.5). The cluster {32,...,45} was found by the CCC in the full data set for all the cluster methods used (Table 3.5) and remained in the consensus classification (Table 3.6).

This distinctiveness relative to the other taxa within the *S. sanctipauli* subcomplex can be explained by reference to the h-plot of the dispersion matrix (Figure 3.1), the original data set (Table 3.2) and the principal components (Table 3.3). Pointing directly at *S. soubrense* 'B' on the h-plot are the inversions 1L-A and 2S-7. The former is a fixed, unique derived character for this species (Post 1986) and the latter is fixed in this species, although it is shared by other members of the subcomplex. Inversions which also have a strong though less direct effect on the distinctiveness of *S. soubrense* 'B' are the inversions 2L-D and 1S-A. 2L-D is fixed in this species but shared by *S. soubrense* except for *Chutes Milo* form. The vector for this inversion on the h-plot is almost coincidental with the first principal axis, so samples along this axis increase in frequency for the inversion. 1S-A, however is orthogonal to the first principal axis, and hence uncorrelated with inversion 2L-D. This inversion is also fixed in *S. soubrense* 'B' but is shared with other members of the subcomplex.

When the B_k methods of cluster analysis were applied to analyse variation within *S. soubrense* 'B', the distortion measure $\Delta_{0.5}$ fell

only slightly as k was increased from one to four (Table 3.7); such a high level of distortion suggests that there is little structure within the data set, either hierarchical in nature or overlapping (Jardine and Sibson 1968). The two-dimensional ordination resulting from a principal components analysis of the dispersion matrix is shown as Figure 3.6, and superimposed on this ordination are the clusters resulting from application of the $B_{k=2}$ method at $h=0.001$ (chosen because of a distinct moat at this level). This three cluster solution shows that the bulk of the data belong to a homogeneous cluster, with two outlying samples. Three inversions are important in this variation, 1S-C, 3L-4, 3L-17. The small angle between the last two on the h -plot reflects their strong linkage. 1S-C was only found at a low frequency in *S. soubrense* 'B' (for which species it is unique), with the maximum frequency for sample 40. Thus the variation along the second principal axis is dominated by this relatively unimportant inversion, demonstrating that there is in fact very little intraspecific variation. 3L-4 and 3L-17 are partially X-linked in *S. soubrense* 'B' (Post 1986), although this information was not included in the analysis. As expected from the h -plot, the frequency of these inversions is lowest in sample 44 and largest in sample 41. This is because sample 41 was a sample of six males and four females, while sample 44 was a sample of all females.

To conclude *S. soubrense* 'B' represents a chromosomally very homogeneous taxon which is very distinct from other members of the subcomplex, a result which supports the conclusions of classical cytotaxonomy (Post 1986). The only intraspecific variation found is likely to be because of random variation and sex ratio differences rather than being due to any systematic geographic or temporal vari-

ation. The considerable number of fixed inversions in this taxon (including inversions shared with other members of the subcomplex, Table 3.2) and the species' restricted geographic distribution suggests that the founding population for this species may have been small (i.e. the origin of the species involved a population bottleneck, Mayr 1970), or that the species is very ancient.

3.4.2 SIMULIUM SOUBRENSE 'CHUTES MILO/BEFFA'

By contrast *S. soubrense* 'Chutes Milo/Beffa' is an ill-defined taxon, both on the ordination (Figure 3.1) and on the dendrograms (Figures 3.2-3.5). Samples 1-4 belong to *S. soubrense* 'Beffa', but this taxon was not uncovered by any of the cluster methods and was not defined in the consensus classification (Table 3.6). This may in part be because of the information contained in the partially Y-linked inversion 2S-6b not being included in the analysis. Samples 5-8 correspond to *S. soubrense* 'Chutes Milo/Typical' and is much better defined, although only the three member cluster {6,7,8} remained in the consensus classification. On the ordination these samples lie close to *S. sanctipauli*, the only major inversion difference between these taxa being the inversion 2L-A which is fixed in *S. sanctipauli* and absent from *S. soubrense* (Post 1986).

Because of the small number of samples, *S. soubrense* 'Beffa' and *S. soubrense* 'Chutes Milo' were not analysed separately from the other samples.

To conclude *S. soubrense* 'Beffa' and *S. soubrense* 'Chutes Milo' are relatively heterogeneous taxa, although the former is more so than the latter. The very small number of samples does not allow for interpretation either in terms of geography or time. The close proximity of these samples to *S. sanctipauli* is of considerable interest,

especially when compared with the large taxonomic distance between either of these taxa and *S. soubrense* 'B'.

3.4.3 SIMULIUM SANCTIPAULI

Simulium sanctipauli (samples 46-66) appears as a heterogeneous taxon both on the ordination (Figure 3.1) and on the dendrograms (Figures 3.2-3.5). The principal inversion distinguishing this species from the rest of the subcomplex is 2L-A (Post 1986), which is a unique fixed derived character for *S. sanctipauli*. On the ordination, much of the variation within *S. sanctipauli* is because of variation in the inversion 1S-A, as is clear from the h-plot and the original data matrix (Table 3.2).

The consensus partition of the whole data set (Table 3.6) resulted in three consistent clusters relative to the total data set {46, ..., 50}, {54, 55, 56, 63, 64} and {57, 59, ..., 62}. Some sample points were not included in the consensus partition because they were inconsistently classified using the different cluster methods. The first cluster corresponds to Djodji form (Chapter two), the second are samples from the Comoe, Maraoue and Baoule rivers, while the last corresponds to samples from the Bandama and Sassandra rivers.

Application of the B_k methods of Jardine and Sibson (1968) to the *S. sanctipauli* data in isolation resulted in quite a marked reduction in the distortion imposed by the resultant classification (Table 3.7). This result implies that the data can more realistically be represented by allowing samples to overlap between clusters, although the high distortion remaining at $B_{k=4}$ showed that allowing for overlap has not entirely removed distortion from the classification.

The two-dimensional ordination resulting from a principal components analysis of the dispersion matrix is shown as Figure 3.7, with

the clusters defined at $h=0.001$ for the method $B_{k=2}$. This level was chosen because of a distinct moat between this level and the next fusion. Seven clusters were identified in the data set, but four of these overlap. The most distinctive cluster corresponds to the samples from the Rivers Maraoue, Comoe and Baoule. These samples differ from other *S. sanctipauli* principally in the inversion 2L-7 which is absent from this cluster but at a high frequency in other clusters (as can be seen from the h-plot and Table 3.2). Also the inversion 3L-B which is absent from most other *S. sanctipauli* is present at a high frequency in this cluster. Despite the diverse geographic origin of the members of this cluster, it is internally homogeneous. The most important practical aspect of this cluster is that four of the samples (55,56,63,64) are the only samples within the *S. sanctipauli* subcomplex resistant to organo-phosphate insecticide (Kurtak and Post 1987).

The samples 46-50 form another relatively tight cluster, although one sample (49) is shared in common with another cluster. These samples correspond to the Djodji form of *S. sanctipauli* defined on classical criteria by Surtees *et al.* (1988). The information contained in the strongly Y-linked inversion 1S-21 was not incorporated in the analysis. If it had been then the cluster would have been more distinct.

The cluster {57,59,60,61,62} has two samples which overlap with other clusters. One sample overlaps with the cluster {57,58} both from the Bandama river, Cote d'Ivoire, but separated by 11 years (Table 3.1), the other cluster is the Djodji cluster. The samples within this cluster are all from the rivers Bandama and Sassandra. and are distinguished by a high frequency for inversion 1S-A.

The cluster {53,58} overlaps with the Bandama cluster and includes the Ghanaian sample from the River Ejisu. The cluster {51,52} is from the River Pra, Ghana, while the geographically isolated cluster {65,66} is from the River Moa in Sierra Leone, the most westerly of the *S. sanctipauli* samples. This last cluster is particularly unusual, lacking inversion 1S-A and sharing three inversions not shared with other *S. sanctipauli*: 1S-F, 1L-B and 1L-C.

To conclude, there is considerable variation within *S. sanctipauli*, at least some of which is likely to involve restriction of gene flow between clusters. The correlation of OP insecticide resistance with one cluster is strong evidence for restricted gene flow between it and the other clusters within *S. sanctipauli*. This feature will be very useful for tracing OP resistance movement. It is unlikely, therefore, that *S. sanctipauli* will remain as a unitary taxon once more information becomes available, and the Djodji form of *S. sanctipauli* defined on classical criteria (Chapter two) will not be the only cytoform within *S. sanctipauli*.

3.4.4 SIMULIUM SOUBRENSE 'MENANKAYA/KONKOURE'

Simulium soubrense 'Menankaya/Konkoure' lies on a broad band in the lower right quadrant of the ordination (Figure 3.1) and shows a considerable degree of chromosomal heterogeneity. The superimposed MST on figure 3.1 reveals that samples 9 and 16, which are close on the ordination are in fact quite distinct, showing that the ordination has introduced some distortion into the data. The h-plot of the dispersion matrix shows that several inversions define the distinction between this taxon and the other members of the subcomplex. Inversion 2L-D, which is coincident with the first principal axis is present in all the samples of *S. soubrense* 'Menankaya/Konkoure' (Table 3.2),

but it varies in frequency from fixation to 0.25. However, this inversion mainly serves to define the right half of the ordination, including *S. soubrense* 'B', but excluding *S. soubrense* 'Chutes Milo/Beffa' and *S. sanctipauli*. The inversion 1L-A defines *S. soubrense* 'B' in the upper right quadrant, while inversion 1S-A which is present throughout the subcomplex is found at a low but variable frequency in *S. soubrense* 'Menankaya/Konkoure'. The inversion 2S-7 which has a strong influence at 30° to the first principal axis varies within *S. soubrense* 'Menankaya/Konkoure', as do the inversions 2L-X and 3L-X.

The different cluster methods all found the same clusters within *S. soubrense* 'Menankaya/Konkoure' (Table 3.5) and these remained in the consensus classification (Table 3.6). There are six clusters, although two of these are singletons (samples 9 and 14).

Applying the B_k methods to *S. soubrense* 'Menankaya/Konkoure' resulted in some reduction in the distortion resulting from the transformation from distance matrix to ultrametric matrix (Table 3.7) but this reduction was not as great as would be expected if the data were sampled from genuinely overlapping clusters (Jardine and Sibson, 1968, found that the measure of distortion fell from 0.528 to 0.146 as k was increased from one to four for overlapping clusters of the annual pearlwort *Sagina apetala*).

Figure 3.8 shows the ordination resulting from a principal components analysis of the dispersion matrix, and reveals the complexity of variation within this taxon. The four cluster and two singletons are marked. The four main clusters are connected by the MST in a sequence forming a horseshoe on the ordination from *S. soubrense* 'Menankaya' (samples 10-13) from Sierra Leone/Guinea, to *S. soubrense*

'Menankaya/Konkoure' (samples 15-23) from the Rivers Tene and Bafing in the Fouta Djallon, Guinea, to *S. soubrense* 'Konkoure' (samples 30,31) from the River Koumba, Guinea, to *S. soubrense* 'Konkoure' (samples 24-29) from the Konkoure and Kakrima rivers in Guinea. A sequence of clusters connected in this way suggests clinal variation with inadequate sampling between clusters (Jardine 1971). The h-plot reveals that several inversions influence the observed pattern of variation. Inversion 2S-7 is found in the Konkoure/Koumba samples, and in the Sierra Leone samples, although its frequency is much lower in the latter. 3L-X also defines the Konkoure/Koumba samples. 3L-2 has an opposite influence to this inversion, being highest in samples 15-23 and lowest in the Konkoure form. The main inversions defining *S. soubrense* 'Menankaya' are 3L-4, 3L-17 (which are linked) and 3L-5. This last inversion is a unique polymorphic derived character defining *S. soubrense* 'Menankaya'.

To conclude *S. soubrense* 'Menankaya/Konkoure' shows a very complex pattern of intraspecific variation involving a considerable number of inversions. Two extreme hypotheses can be established to explain these results. The first is that the data have been sampled inadequately from continuous clinal variation. The sequential pattern shown on the ordination supports this hypothesis. The second hypothesis is that one or all of the clusters uncovered represents distinct forms which in sympatry would not introgress. The distinctiveness and internal homogeneity of the derived clusters supports this hypothesis. Based on the available data it is not possible to choose between these alternatives. The most likely explanation is a combination of the two i.e. that there is clinal var-

iation within *S. soubrense* 'Konkoure', but that *S. soubrense* 'Menankaya' is a form distinct from this.

3.5 CONCLUSIONS

The multivariate statistical analyses presented in this chapter usually support the findings of classical cytotaxonomy, but also identify features of the data which classical methods have overlooked such as the distinctiveness of the OP resistant *S. sanctipauli* flies, and the complex variation within *S. soubrense* 'Menankaya/Konkoure'. These results could now be used predictively to identify and trace OP resistance.

However, the methods could undoubtedly be improved on to make them more sensitive and powerful in the future. One important feature of blackfly cytotaxonomy which has not been incorporated in this analysis is the information contained in the sex-linkage of inversions. The importance of sex-linkage in blackflies has been stated before (e.g. Post 1982). To use this information, the sexes could be treated separately within each sample, although this would require larger sample sizes. To exploit the interpretative power of the multivariate methods, systematic sampling would be an improvement over the *ad hoc* sampling of the data in this analysis. Finally, these same methods could be used on individuals within samples, to identify sympatric taxa.

Table 3.1

List of larval samples of the *S. sanctipauli* subcomplex.

| Sample No. and origin ¹ | River | Site | Coordinates ² (Lat./Long) | Date | Deter- -ined by ³ | <i>a priori</i> taxon |
|------------------------------------|-------------|---------------|---|----------|------------------------------------|---|
| 1 N | Oshun | Ede | 07°44' 05°34'E | 14.07.82 | A | <i>Simulium</i> <i>soubrense</i> Beffa |
| 2 N | Ogun | Eruwa | 07°25' 04°29'E | 15.07.82 | A | |
| 3 T | Mono | Tetetou | 07°02' 01°32'E | 20.07.80 | A | |
| 4 T | Mono | Avegode | 06°48' 01°36'E | 14.11.85 | A | |
| 5 C | Leraba | Leraba Bridge | 10°10' 05°04'W | 10.06.71 | A | <i>Simulium</i> <i>soubrense</i> Chutes Milo |
| 6 C | Cestos | Darlu | - | 06.05.71 | A | |
| 7 C | Niandan | Rapide Pampan | 10°01' 09°41'W | 21.12.85 | A | |
| 8 C | Niandan | Boria | 09°28' 09°57'W | 14.02.86 | A | |
| 9 G | Milo | Tiekoradu | 09°00' 08°57'W | - | A | <i>Simulium</i> <i>soubrense</i> Menankaya |
| 10 S | Sewa | Njaime-Sewafe | 08°52' 11°13'W | 30.11.80 | A | |
| 11 S | Sewa | Babawahun | 07°59' 11°20'W | 15.12.81 | A | |
| 12 S | Seli | Badala | 09°19' 11°32'W | 18.12.81 | A | |
| 13 S | Rokel | Bumbuna | 09°03' 11°44'W | 29.11.81 | A | |
| 14 S | Mongo | Musaia | 09°46' 11°28'W | 17.12.81 | A | |
| 15 G | Bafing | Sokotoro | 10°38' 11°45'W | 13.11.86 | A | <i>Simulium</i> <i>soubrense</i> Konkoure/ Menankaya |
| 16 G | Bafing | Lago | 10°51' 11°36'W | 22.12.85 | A | |
| 17 G | Bafing | Nduria | 10°45' 11°45'W | 13.02.86 | A | |
| 18 G | Bafing | Yagui | 11°34' 10°52'W | 22.11.86 | A | |
| 19 G | Tene | below bridge | 11°01' 11°49'W | 13.02.86 | A | |
| 20 G | Tene | Dankolo | 11°01' 11°58'W | 22.11.86 | A | |
| 21 G | Tene | above Chutes | 11°01' 11°49'W | 22.11.86 | A | |
| 22 G | Tene | above Bafing | 11°07' 11°35'W | 22.11.86 | A | |
| 23 G | Bafing | Koukotamba | 11°13' 11°19'W | 22.11.86 | A | |
| 24 G | Konkoure | Ganiya | 10°29' 12°59'W | 13.02.86 | A | |
| 25 G | Konkoure | Soukia | 10°25' 13°10'W | 12.11.86 | A | |
| 26 G | Konkoure | Bakere | 10°31' 13°10'W | 12.11.86 | A | |
| 27 G | Konkoure | Kanhan | 10°28' 12°47'W | 12.11.86 | A | |
| 28 G | Kakrima | Bougoula | 10°34' 12°58'W | 12.11.86 | A | |
| 29 G | Kakrima | Kaffima | 10°50' 12°57'W | 12.11.86 | A | |
| 30 G | Koumba | Sidipo | 11°43' 12°56'W | 11.11.86 | A | |
| 31 G | Koumba | Kokou | 11°42' 12°54'W | 11.11.86 | A | |
| 32 S | Moa | Tiwai Island | 07°32' 11°22'W | 16.09.83 | A | <i>Simulium</i> <i>soubrense</i> 'B' |
| 33 S | Waanje | Bandajuma | 07°34' 11°39'W | 21.06.83 | A | |
| 34 S | Sewa | Wubunge | 07°48' 11°48'W | 12.06.83 | A | |
| 35 S | Sewa | Mofwe | 07°40' 11°58'W | 08.12.81 | A | |
| 36 S | Tabe | Gbalima | 08°06' 11°51'W | 08.12.80 | A | |
| 37 S | Teye | Mongeri | 08°19' 11°44'W | 11.08.83 | A | |
| 38 S | Taia | Mogbamu | 08°01' 12°07'W | 09.12.80 | A | |
| 39 S | Taia | Mogbamu | 08°01' 12°07'W | 10.06.83 | A | |
| 40 S | Gbangaia | Mokasi | 07°59' 12°25'W | 07.12.81 | A | |
| 41 S | Rokel | Katik | 08°39' 12°30'W | 10.12.80 | A | |
| 42 S | Bankasoka | Port Loko | 08°46' 12°47'W | 12.12.80 | A | |
| 43 S | Gt Scarcies | Kanka | 09°43' 12°27'W | 02.12.81 | A | |

| | | | | | | | | |
|----|----|-----------|---------------|--------|---------|----------|---|---------------------------------------|
| 44 | G | Kolente | Malea | 10°41' | 12°37'W | 12.02.86 | A | |
| 45 | G | Kolente | Kolente | 10°04' | 12°38'W | 13.11.86 | A | |
| 46 | T | Gban-houa | Djodji | 07°40' | 00°35'E | 15.10.84 | B | <i>Simulium sanctipauli</i> Djodji |
| 47 | T | Gban-houa | Djodji | 07°40' | 00°35'E | 15.03.85 | B | |
| 48 | T | Gban-houa | Djodji | 07°40' | 00°35'E | 21.03.85 | B | |
| 49 | T | Gban-houa | Djodji | 07°40' | 00°35'E | 26.03.85 | B | |
| 50 | T | Gban-houa | Djodji | 07°40' | 00°35'E | 29.03.85 | B | |
| 51 | Gh | Pra | Hemang | 05°11' | 01°32'W | 08.07.80 | A | <i>Simulium sanctipauli</i> |
| 52 | Gh | Pra | Hemang | 05°11' | 01°32'W | 17.07.82 | A | |
| 53 | Gh | Ofin | Ejisu | 05°57' | 01°42'W | 25.01.86 | A | |
| 54 | C | Comoe | Mbaso | 06°20' | 03°30'W | 27.07.80 | A | |
| 55 | C | Comoe | Amouakro | - | - | 14.02.85 | A | |
| 56 | C | Maraoue | Danangoro | 07°10' | 05°56'W | 10.08.82 | A | |
| 57 | C | Bandama | Tiassale | 05°53' | 04°49'W | 08.07.82 | A | |
| 58 | C | Bandama | Ahoauti | 06°07' | 04°57'W | 23.06.71 | A | |
| 59 | C | Sassandra | Soubre | 05°47' | 06°37'W | 10.07.82 | A | |
| 60 | C | Sassandra | Soubre | 05°47' | 06°37'W | 26.10.84 | B | |
| 61 | C | Sassandra | Koperagui | - | - | 22.01.85 | A | |
| 62 | C | Sassandra | Chutes Nawa | 05°47' | 06°37'W | 06.09.84 | A | |
| 63 | M | Bouale | Madina Diasso | 10°40' | 07°40'W | 26.01.86 | A | |
| 64 | M | Bouale | Konigbougeula | 10°45' | 07°46'W | 27.01.86 | A | |
| 65 | S | Moa | Maloma | 08°00' | 10°50'W | 06.12.80 | A | |
| 66 | S | Moa | Tiwai Island | 07°33' | 11°22'W | 16.09.83 | A | |

¹N=Nigeria, C=Côte d'Ivoire, T=Togo, Gh=Ghana, G=Guinea, S=Sierra Leone, M=Mali

²All Latitudes are North.

³A=Determined by Dr. R.J. Post, B=Determined by D.P. Surtees

Table 3.2

Polymorphic Inversions within the *S. sanctipauli* subcomplex

Chromosome one inversions

| Sample Number | Sample Size | 1L-B | 1L-C | 1L-D | 1L-U | 1L-G | 1L-N | 1L-X | 1L-A | 1L-R | 1L-T |
|---------------|-------------|-------|-------|-------|------|------|------|-------|------|------|-------|
| 1 | 5 | 0.083 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2 | 14 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 3 | 12 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 4 | 10 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 26 | 0.115 | 0.115 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 11 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 7 | 5 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 8 | 8 | 0.333 | 0.333 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 9 | 8 | 0.688 | 0.75 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10 | 13 | 0.038 | 0.038 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 11 | 8 | 0 | 0 | 0.063 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 12 | 7 | 0.214 | 0.214 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 13 | 11 | 0.045 | 0.045 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 14 | 25 | 0.54 | 0.54 | 0 | 0 | 0.02 | 0 | 0 | 0 | 0 | 0 |
| 15 | 21 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 16 | 12 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 17 | 11 | 0.05 | 0.05 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 18 | 10 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 19 | 8 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 20 | 23 | 0 | 0 | 0 | 0 | 0 | 0 | 0.023 | 0 | 0 | 0 |
| 21 | 24 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 22 | 11 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 23 | 19 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 24 | 17 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 25 | 29 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 26 | 17 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 27 | 16 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 28 | 20 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 29 | 31 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 30 | 23 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 31 | 32 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 32 | 28 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 33 | 24 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 34 | 39 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 35 | 15 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 36 | 12 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 37 | 22 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 38 | 22 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 39 | 40 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 40 | 8 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0.013 |
| 41 | 10 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 42 | 25 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 43 | 42 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 44 | 12 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 45 | 14 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 46 | 29 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

| | | | | | | | | | | | |
|----|----|-------|-------|---|-------|---|-------|---|-------|-------|---|
| 47 | 23 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 48 | 23 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 49 | 32 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 50 | 21 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 51 | 16 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 52 | 30 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 53 | 7 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 54 | 45 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 55 | 7 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 56 | 14 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 57 | 43 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.012 | 0 |
| 58 | 21 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 59 | 18 | 0 | 0 | 0 | 0 | 0 | 0.028 | 0 | 0 | 0 | 0 |
| 60 | 24 | 0 | 0 | 0 | 0.021 | 0 | 0 | 0 | 0 | 0 | 0 |
| 61 | 5 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 62 | 6 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 63 | 12 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 64 | 10 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 65 | 8 | 0.5 | 0.5 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 66 | 21 | 0.262 | 0.262 | 0 | 0 | 0 | 0 | 0 | 0.024 | 0 | 0 |

Table 3.2 (continued)

Chromosome one inversions

| Sample Number | 1S-A | 1S-G. | 1S-J | 1S-F | 1S-B | 1S-X | 1S-N | 1S-C | 1S-M | 1S-P | 1S-21 |
|---------------|-------|-------|-------|------|-------|-------|-------|-------|-------|------|-------|
| 1 | 0.5 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2 | 0.821 | 0.179 | 0.036 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 3 | 1 | 0.818 | 0.136 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 4 | 1 | 0.05 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 7 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 9 | 0.063 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 11 | 0.188 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 12 | 0.357 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 13 | 0.182 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 14 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 15 | 0 | 0 | 0 | 0 | 0.105 | 0 | 0 | 0 | 0 | 0 | 0 |
| 16 | 0.417 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 17 | 0.1 | 0 | 0.05 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 18 | 0.05 | 0 | 0 | 0 | 0.05 | 0 | 0 | 0 | 0 | 0 | 0 |
| 19 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 20 | 0 | 0 | 0 | 0 | 0.095 | 0 | 0 | 0 | 0 | 0 | 0 |
| 21 | 0.026 | 0 | 0.026 | 0 | 0.132 | 0 | 0 | 0 | 0 | 0 | 0 |
| 22 | 0 | 0 | 0 | 0 | 0.056 | 0 | 0 | 0 | 0 | 0 | 0 |
| 23 | 0 | 0 | 0 | 0 | 0.033 | 0 | 0 | 0 | 0 | 0 | 0 |
| 24 | 0.036 | 0 | 0.071 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 25 | 0.019 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 26 | 0.033 | 0 | 0 | 0 | 0.033 | 0 | 0 | 0 | 0 | 0 | 0 |
| 27 | 0.125 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 28 | 0 | 0 | 0 | 0 | 0.118 | 0.059 | 0 | 0 | 0 | 0 | 0 |
| 29 | 0.054 | 0 | 0 | 0 | 0.125 | 0.036 | 0 | 0 | 0 | 0 | 0 |
| 30 | 0.306 | 0 | 0 | 0 | 0.306 | 0 | 0.028 | 0 | 0 | 0 | 0 |
| 31 | 0.466 | 0 | 0 | 0 | 0.293 | 0 | 0 | 0 | 0 | 0 | 0 |
| 32 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0.071 | 0 | 0 | 0 |
| 33 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0.083 | 0 | 0 | 0 |
| 34 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0.205 | 0 | 0 | 0 |
| 35 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0.2 | 0 | 0 | 0 |
| 36 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0.083 | 0 | 0 | 0 |
| 37 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0.159 | 0 | 0 | 0 |
| 38 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0.136 | 0 | 0 | 0 |
| 39 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0.167 | 0 | 0 | 0 |
| 40 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0.188 | 0 | 0 | 0 |
| 41 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0.2 | 0 | 0 | 0 |
| 42 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0.04 | 0 | 0 | 0 |
| 43 | 1 | 0 | 0 | 0 | 0 | 0 | 0.012 | 0.036 | 0.012 | 0 | 0 |
| 44 | 0.875 | 0 | 0.125 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 45 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 46 | 0.446 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.214 |
| 47 | 0.543 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.261 |
| 48 | 0.321 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.184 |

| | | | | | | | | | | | |
|----|-------|---|---|-------|-------|---|---|---|---|-------|-------|
| 49 | 0.613 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.016 | 0.145 |
| 50 | 0.523 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.176 |
| 51 | 0.344 | 0 | 0 | 0 | 0.063 | 0 | 0 | 0 | 0 | 0 | 0 |
| 52 | 0.133 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.017 |
| 53 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.071 |
| 54 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 55 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 56 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 57 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 58 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 59 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 60 | 0.979 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 61 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 62 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 63 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 65 | 0 | 0 | 0 | 0.563 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 66 | 0 | 0 | 0 | 0.31 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

Table 3.2 (continued)

Chromosome two inversions

| Sample Number | 2L-A | 2L-7 | 2L-D | 2L-X | 2L-S | 2L-W | 2L-B | 2S-7 | 2S-6b |
|---------------|------|-------|-------|-------|------|-------|-------|-------|-------|
| 1 | 0 | 0.9 | 0 | 0 | 0 | 0 | 0 | 0 | 0.167 |
| 2 | 0 | 0.714 | 0 | 0 | 0 | 0 | 0 | 0 | 0.25 |
| 3 | 0 | 0.091 | 0.364 | 0 | 0 | 0 | 0 | 0 | 0.125 |
| 4 | 0 | 0.05 | 0.45 | 0 | 0 | 0 | 0 | 0 | 0.05 |
| 5 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 7 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 8 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 9 | 0 | 0.938 | 0.938 | 0 | 0 | 0 | 0 | 0.625 | 0 |
| 10 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 0.154 | 0 |
| 11 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 0.25 | 0 |
| 12 | 0 | 0.857 | 1 | 0 | 0 | 0 | 0 | 0.071 | 0 |
| 13 | 0 | 0.864 | 1 | 0 | 0 | 0 | 0 | 0.227 | 0 |
| 14 | 0 | 0.22 | 1 | 0 | 0 | 0 | 0 | 0.02 | 0 |
| 15 | 0 | 0.762 | 0.762 | 0.238 | 0 | 0 | 0 | 0 | 0 |
| 16 | 0 | 0.833 | 0.833 | 0.167 | 0 | 0 | 0 | 0 | 0 |
| 17 | 0 | 0.636 | 0.636 | 0.364 | 0 | 0 | 0 | 0 | 0 |
| 18 | 0 | 0.85 | 0.85 | 0.15 | 0 | 0 | 0 | 0 | 0 |
| 19 | 0 | 0.75 | 0.75 | 0.25 | 0 | 0 | 0 | 0 | 0 |
| 20 | 0 | 0.826 | 0.826 | 0.174 | 0 | 0 | 0 | 0 | 0 |
| 21 | 0 | 0.833 | 0.833 | 0.167 | 0 | 0 | 0 | 0 | 0 |
| 22 | 0 | 0.909 | 0.909 | 0.091 | 0 | 0 | 0 | 0 | 0 |
| 23 | 0 | 0.895 | 0.895 | 0.105 | 0 | 0 | 0 | 0 | 0 |
| 24 | 0 | 0.794 | 0.794 | 0.206 | 0 | 0 | 0 | 0.912 | 0 |
| 25 | 0 | 0.821 | 0.821 | 0.179 | 0 | 0 | 0 | 0.931 | 0 |
| 26 | 0 | 0.875 | 0.875 | 0.125 | 0 | 0 | 0 | 1 | 0 |
| 27 | 0 | 0.7 | 0.7 | 0.3 | 0 | 0 | 0 | 0.875 | 0 |
| 28 | 0 | 0.65 | 0.65 | 0.35 | 0 | 0 | 0 | 1 | 0 |
| 29 | 0 | 0.683 | 0.683 | 0.317 | 0 | 0 | 0 | 0.774 | 0 |
| 30 | 0 | 0.25 | 0.25 | 0.75 | 0 | 0 | 0 | 0 | 0 |
| 31 | 0 | 0.328 | 0.328 | 0.672 | 0 | 0 | 0 | 0 | 0 |
| 32 | 0 | 1 | 1 | 0 | 0 | 0.018 | 0 | 1 | 0 |
| 33 | 0 | 1 | 1 | 0 | 0 | 0 | 0.022 | 1 | 0 |
| 34 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 |
| 35 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 |
| 36 | 0 | 1 | 1 | 0 | 0 | 0 | 0.042 | 1 | 0 |
| 37 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 |
| 38 | 0 | 1 | 1 | 0 | 0 | 0 | 0.023 | 1 | 0 |
| 39 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 |
| 40 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 |
| 41 | 0 | 1 | 1 | 0 | 0 | 0 | 0.05 | 1 | 0 |
| 42 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 |
| 43 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 |
| 44 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 |
| 45 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 |
| 46 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 47 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 48 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

| | | | | | | | | | |
|----|---|-------|---|---|-------|---|---|---|---|
| 49 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 50 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 51 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 52 | 1 | 0.983 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 53 | 1 | 0.714 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 54 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 55 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 56 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 57 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 58 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 59 | 1 | 0.833 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 60 | 1 | 0.854 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 61 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 62 | 1 | 0.917 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 63 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 65 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 66 | 1 | 0.976 | 0 | 0 | 0.024 | 0 | 0 | 0 | 0 |

Table 3.2 (continued)

Chromosome three inversions

| Sample Number | 3L-A | 3L-24 | 3L-B | 3L-25 | 3L-26 | 3L-5 | 3L-E | 3L-I | 3L-G | 3L-Y |
|---------------|------|-------|-------|-------|-------|-------|-------|-------|-------|-------|
| 1 | 0 | 0.25 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2 | 0 | 0.357 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 3 | 0 | 0.75 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 4 | 0 | 0.75 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 0 | 0.583 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 0 | 0 | 0 | 0.2 | 0.15 | 0 | 0 | 0 | 0 | 0 |
| 7 | 0 | 0 | 0.167 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 8 | 0 | 0 | 0.125 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 9 | 0 | 0 | 0 | 0 | 0 | 0.313 | 0 | 0 | 0 | 0 |
| 10 | 0 | 0 | 0 | 0 | 0 | 0.731 | 0.038 | 0.038 | 0.077 | 0 |
| 11 | 0 | 0 | 0 | 0 | 0 | 0.625 | 0.063 | 0 | 0 | 0 |
| 12 | 0 | 0 | 0 | 0 | 0 | 0.714 | 0 | 0 | 0 | 0 |
| 13 | 0 | 0 | 0 | 0 | 0 | 0.72 | 0 | 0 | 0 | 0 |
| 14 | 0 | 0 | 0 | 0 | 0 | 0.72 | 0 | 0 | 0 | 0 |
| 15 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 16 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 17 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.111 |
| 18 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 19 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 20 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 21 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 22 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 23 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 24 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 25 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.069 | 0 | 0 |
| 26 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.029 | 0 | 0 |
| 27 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.033 | 0 | 0 |
| 28 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 29 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 30 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 31 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 32 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 33 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 34 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 35 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 36 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 37 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 38 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 39 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 40 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 41 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 42 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.02 | 0 | 0 |
| 43 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.024 | 0 | 0 |
| 44 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 45 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 46 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 47 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 48 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

| | | | | | | | | | | |
|----|-------|-------|-------|---|---|---|---|-------|---|---|
| 49 | 0 | 0.016 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 50 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 51 | 0 | 0 | 0.563 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 52 | 0 | 0 | 0.7 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 53 | 0 | 0 | 0.714 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 54 | 0.022 | 0 | 0.911 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 55 | 0 | 0 | 0.857 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 56 | 0 | 0 | 0.893 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 57 | 0 | 0 | 0.267 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 58 | 0 | 0 | 0.643 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 59 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0.056 | 0 | 0 |
| 60 | 0.021 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 61 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 62 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 63 | 0 | 0 | 0.917 | 0 | 0 | 0 | 0 | 0.042 | 0 | 0 |
| 64 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 65 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 66 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

Table 3.2 (continued)

Chromosome three inversions (continued)

| Sample Number | 3L-X | 3L-2 | 3L-4 | 3L-17 | 3L-D | 3L-K | 3L-M |
|---------------|-------|-------|-------|-------|-------|-------|------|
| 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 2 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 3 | 0 | 1 | 0.958 | 0.958 | 0 | 0 | 0 |
| 4 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 5 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 6 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 7 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 8 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 9 | 0 | 1 | 0.75 | 0.75 | 0 | 0 | 0 |
| 10 | 0 | 1 | 0.962 | 0.962 | 0 | 0 | 0 |
| 11 | 0 | 1 | 0.937 | 0.937 | 0 | 0 | 0 |
| 12 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 13 | 0 | 1 | 0.94 | 0.94 | 0 | 0 | 0 |
| 14 | 0 | 1 | 0.94 | 0.94 | 0 | 0 | 0 |
| 15 | 0 | 1 | 0.286 | 0.286 | 0 | 0 | 0 |
| 16 | 0 | 1 | 0.333 | 0.333 | 0 | 0 | 0 |
| 17 | 0 | 1 | 0.389 | 0.389 | 0 | 0 | 0 |
| 18 | 0.05 | 0.95 | 0.1 | 0.1 | 0 | 0 | 0 |
| 19 | 0.313 | 1 | 0.188 | 0.188 | 0 | 0 | 0 |
| 20 | 0 | 1 | 0.275 | 0.275 | 0 | 0 | 0 |
| 21 | 0 | 1 | 0.31 | 0.31 | 0 | 0 | 0 |
| 22 | 0 | 1 | 0.45 | 0.45 | 0 | 0 | 0 |
| 23 | 0 | 1 | 0.313 | 0.313 | 0 | 0 | 0 |
| 24 | 0.813 | 0.125 | 0.125 | 0.125 | 0 | 0 | 0 |
| 25 | 0.741 | 0.258 | 0.258 | 0.258 | 0 | 0 | 0 |
| 26 | 0.765 | 0.118 | 0.118 | 0.118 | 0 | 0 | 0 |
| 27 | 0.733 | 0.267 | 0.267 | 0.267 | 0 | 0 | 0 |
| 28 | 0.7 | 0.3 | 0.3 | 0.3 | 0 | 0 | 0 |
| 29 | 0.839 | 0.161 | 0.143 | 0.143 | 0 | 0 | 0 |
| 30 | 0.636 | 0.364 | 0.227 | 0.227 | 0 | 0 | 0 |
| 31 | 0.672 | 0.034 | 0.293 | 0.293 | 0 | 0 | 0 |
| 32 | 0 | 1 | 0.42 | 0.38 | 0.054 | 0 | 0 |
| 33 | 0 | 1 | 0.286 | 0.238 | 0.022 | 0 | 0 |
| 34 | 0 | 1 | 0.357 | 0.3 | 0.013 | 0 | 0 |
| 35 | 0 | 1 | 0.357 | 0.321 | 0.033 | 0 | 0 |
| 36 | 0 | 1 | 0.278 | 0.278 | 0.083 | 0 | 0 |
| 37 | 0 | 1 | 0.333 | 0.357 | 0.045 | 0 | 0 |
| 38 | 0 | 1 | 0.406 | 0.437 | 0.068 | 0 | 0 |
| 39 | 0 | 1 | 0.25 | 0.187 | 0.013 | 0 | 0 |
| 40 | 0 | 1 | 0.4 | 0.4 | 0.188 | 0 | 0 |
| 41 | 0 | 1 | 0.55 | 0.55 | 0.1 | 0 | 0 |
| 42 | 0 | 1 | 0.4 | 0.38 | 0 | 0 | 0 |
| 43 | 0 | 1 | 0.342 | 0.329 | 0 | 0.012 | 0 |
| 44 | 0 | 1 | 0.125 | 0.125 | 0 | 0 | 0 |
| 45 | 0 | 1 | 0.423 | 0.423 | 0 | 0 | 0 |
| 46 | 0 | 1 | 0.982 | 0.982 | 0 | 0 | 0 |
| 47 | 0 | 1 | 0.935 | 0.935 | 0 | 0 | 0 |
| 48 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |

| | | | | | | | |
|----|---|---|-------|-------|---|---|-------|
| 49 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 50 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 51 | 0 | 1 | 0.937 | 0.937 | 0 | 0 | 0 |
| 52 | 0 | 1 | 1 | 1 | 0 | 0 | 0.017 |
| 53 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 54 | 0 | 1 | 0.911 | 0.911 | 0 | 0 | 0 |
| 55 | 0 | 1 | 0.857 | 0.857 | 0 | 0 | 0 |
| 56 | 0 | 1 | 0.893 | 0.893 | 0 | 0 | 0 |
| 57 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 58 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 59 | 0 | 1 | 0.944 | 0.944 | 0 | 0 | 0 |
| 60 | 0 | 1 | 0.958 | 0.958 | 0 | 0 | 0 |
| 61 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 62 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 63 | 0 | 1 | 0.917 | 0.917 | 0 | 0 | 0 |
| 64 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 65 | 0 | 1 | 1 | 1 | 0 | 0 | 0 |
| 66 | 0 | 1 | 0.976 | 0.976 | 0 | 0 | 0 |

Table 3.3

First Two Principal Components of the Covariance Matrix

| Inversion | PC I | PC II |
|-----------|-----------|-----------|
| 1L-B | -0.016575 | -0.026580 |
| 1L-C | -0.015252 | -0.027132 |
| 1L-D | 0.000195 | -0.000967 |
| 1L-U | -0.000391 | 0.000290 |
| 1L-G | -0.000063 | 0.000068 |
| 1L-N | -0.000520 | 0.000399 |
| 1L-X | 0.000189 | -0.000742 |
| 1L-A | 0.313813 | 0.502275 |
| 1L-R | -0.000236 | 0.000197 |
| 1L-T | 0.000314 | 0.000453 |
| 1S-A | -0.080477 | 0.662220 |
| 1S-G.H | -0.010888 | 0.006406 |
| 1S-J | 0.003284 | -0.000580 |
| 1S-F | -0.015212 | -0.014282 |
| 1S-B | 0.010259 | -0.052903 |
| 1S-X | 0.001652 | -0.003870 |
| 1S-N | 0.000387 | -0.000829 |
| 1S-C | 0.034813 | 0.057303 |
| 1S-M | 0.000270 | 0.000430 |
| 1S-P | -0.000291 | 0.000056 |
| 1S-21 | -0.019286 | 0.000575 |
| 2L-A | -0.420747 | 0.146415 |
| 2L-7 | 0.170603 | 0.046396 |
| 2L-D | 0.464544 | -0.001129 |
| 2L-X | 0.046113 | -0.174524 |
| 2L-S | -0.000409 | -0.000378 |
| 2L-W | 0.000390 | 0.000661 |
| 2L-B | 0.002974 | 0.005057 |
| 2S-7 | 0.425564 | 0.260556 |
| 2S-6b | -0.006319 | 0.000757 |
| 3L-A | -0.000921 | 0.000646 |
| 3L-24 | -0.030320 | 0.013454 |
| 3L-B | -0.173934 | 0.098074 |
| 3L-25 | -0.002764 | 0.001285 |
| 3L-26 | -0.002073 | 0.000964 |
| 3L-5 | 0.005339 | -0.049871 |
| 3L-E | 0.000280 | -0.001812 |
| 3L-I | 0.001531 | -0.002783 |
| 3L-G | 0.000172 | -0.001713 |
| 3L-Y | 0.000472 | -0.003300 |
| 3L-X | 0.096833 | -0.255164 |
| 3L-2 | -0.099292 | 0.263403 |
| 3L-4 | -0.341389 | 0.130582 |
| 3L-17 | -0.346537 | 0.122688 |
| 3L-D | 0.013501 | 0.022850 |
| 3L-K | 0.000270 | 0.000430 |
| 3L-M | -0.000337 | -0.000135 |

Table 3.4

Distortion measures for each of the hierarchical cluster methods, and the partition level suggested by the cubic clustering criterion for all methods.

| Cluster Method | Distortion Measure | | | | CCC Parti- ion |
|-------------------------|--------------------|-------------|-------------|-------------|----------------------|
| | $r(\text{cop})/$ | $\Delta(0)$ | $\Delta(1)$ | $\Delta(2)$ | |
| Single Linkage | 0.7761 | 0.807 | 0.501 | 0.5412 | 12 |
| Complete Linkage | 0.7898 | 0.8811 | 0.8384 | 0.8355 | 13 |
| Group Average | 0.8428 | 0.4415 | 0.2588 | 0.2119 | 14 |
| Ward's Method | 0.7968 | - | - | - | 14 |
| Iterative Relocation | n.a | n.a | n.a | n.a | 13 |

Table 3.5

Classification results for the five cluster methods on the full data set; partition level set by the cubic clustering criterion; Numbers refer to the samples listed in Table 3.1.

a). Single linkage 12 cluster solution:

| Cluster Number | Sample Number | <i>a priori</i> taxon |
|----------------|---------------------|-------------------------|
| Cluster 1 | 1 2 | Beffa |
| Cluster 2 | 3 | Beffa |
| Cluster 3 | 4, ..., 8 | Beffa/Chutes Milo |
| Cluster 4 | 9 | Menankaya |
| Cluster 5 | 10, ..., 13 | Menankaya |
| Cluster 6 | 14 | Menankaya |
| Cluster 7 | 15, ..., 23 | Konkoure/Menankaya |
| Cluster 8 | 24, ..., 29 | Konkoure/Menankaya |
| Cluster 9 | 30, 31 | Konkoure/Menankaya |
| Cluster 10 | 32, ..., 45 | <i>S. soubrense</i> 'B' |
| Cluster 11 | 46, ..., 53 | <i>S. sanctipauli</i> |
| | 57, ..., 62, 65, 66 | <i>S. sanctipauli</i> |
| Cluster 12 | 54, 55, 56, 63, 64 | <i>S. sanctipauli</i> |

b). Furthest neighbour 13 cluster solution:

| Cluster Number | Sample Number | <i>a priori</i> taxon |
|----------------|--------------------|-------------------------|
| Cluster 1 | 1 2 | Beffa |
| Cluster 2 | 3, 4, 5 | Beffa |
| Cluster 3 | 6, 7, 8 | Chutes Milo |
| Cluster 4 | 9 | Menankaya |
| Cluster 5 | 10, ..., 13 | Menankaya |
| Cluster 6 | 14 | Menankaya |
| Cluster 7 | 15, ..., 23 | Konkoure/Menankaya |
| Cluster 8 | 24, ..., 29 | Konkoure/Menankaya |
| Cluster 9 | 30, 31 | Konkoure/Menankaya |
| Cluster 10 | 32, ..., 45 | <i>S. soubrense</i> 'B' |
| Cluster 11 | 46, ..., 53 | <i>S. sanctipauli</i> |
| | 57, ..., 62 | |
| Cluster 12 | 54, 55, 56, 63, 64 | <i>S. sanctipauli</i> |
| Cluster 13 | 65, 66 | <i>S. sanctipauli</i> |

c). Group Average 14 cluster solution:

| Cluster Number | Sample Number | <i>a priori</i> taxon |
|----------------|------------------------------|-------------------------|
| Cluster 1 | 1 2 | Beffa |
| Cluster 2 | 3 | Beffa |
| Cluster 3 | 4, ..., 8 | Beffa/Chutes Milo |
| Cluster 4 | 9 | Menankaya |
| Cluster 5 | 10, ..., 13 | Menankaya |
| Cluster 6 | 14 | Menankaya |
| Cluster 7 | 15, ..., 23 | Konkoure/Menankaya |
| Cluster 8 | 24, ..., 29 | Konkoure/Menankaya |
| Cluster 9 | 30, 31 | Konkoure/Menankaya |
| Cluster 10 | 32, ..., 45 | <i>S. soubrense</i> 'B' |
| Cluster 11 | 46, ..., 50, 53, 57, ..., 62 | <i>S. sanctipauli</i> |
| Cluster 12 | 51, 52 | <i>S. sanctipauli</i> |
| Cluster 13 | 54, 55, 56, 63, 64 | <i>S. sanctipauli</i> |
| Cluster 14 | 65, 66 | <i>S. sanctipauli</i> |

d). Ward's Method 14 cluster solution:

| Cluster Number | Sample Number | <i>a priori</i> taxon |
|----------------|--------------------|-------------------------|
| Cluster 1 | 1 2 | Beffa |
| Cluster 2 | 3, 4, 5 | Beffa/Chutes Milo |
| Cluster 3 | 6, 7, 8 | Chutes Milo |
| Cluster 4 | 9 | Menankaya |
| Cluster 5 | 10, ..., 13 | Menankaya |
| Cluster 6 | 14 | Menankaya |
| Cluster 7 | 15, ..., 23 | Konkoure/Menankaya |
| Cluster 8 | 24, ..., 29 | Konkoure/Menankaya |
| Cluster 9 | 30, 31 | Konkoure/Menankaya |
| Cluster 10 | 32, ..., 45 | <i>S. soubrense</i> 'B' |
| Cluster 11 | 46, ..., 52 | <i>S. sanctipauli</i> |
| Cluster 12 | 53, 57, ..., 62 | <i>S. sanctipauli</i> |
| Cluster 13 | 54, 55, 56, 63, 64 | <i>S. sanctipauli</i> |
| Cluster 14 | 65, 66 | <i>S. sanctipauli</i> |

e). Iterative Relocation 13 cluster solution:

| Cluster Number | Sample Number | <i>a priori</i> taxon |
|----------------|-----------------|-------------------------|
| Cluster 1 | 1 2 | Beffa |
| Cluster 2 | 3,4,5 | Beffa/Chutes Milo |
| Cluster 3 | 6,7,8 | Chutes Milo |
| Cluster 4 | 9 | Menankaya |
| Cluster 5 | 10,...,13 | Menankaya |
| Cluster 6 | 14 | Menankaya |
| Cluster 7 | 15,...,23 | Konkoure/Menankaya |
| Cluster 8 | 24,...,29 | Konkoure/Menankaya |
| Cluster 9 | 30,31 | Konkoure/Menankaya |
| Cluster 10 | 32,...,45 | <i>S. soubrense</i> 'B' |
| Cluster 11 | 46,...,52 | <i>S. sanctipauli</i> |
| | 57,...,62 | |
| Cluster 12 | 53,...,56,63,64 | <i>S. sanctipauli</i> |
| Cluster 13 | 65,66 | <i>S. sanctipauli</i> |

Table 3.6

Consensus classification¹

| Cluster Number | Sample Number | <i>a priori</i> taxon |
|----------------|-----------------|--------------------------------|
| Cluster 1 | 1,2 | Beffa |
| Cluster 2 | 3 | Beffa |
| Cluster 3 | 6,7,8 | Chutes Milo |
| Cluster 4 | 9 | Menankaya |
| Cluster 5 | 10,...,13 | Menankaya |
| Cluster 6 | 14 | Menankaya |
| Cluster 7 | 15,...,23 | Konkoure/Menankaya |
| Cluster 8 | 24,...,29 | Konkoure/Menankaya |
| Cluster 9 | 30,31 | Konkoure/Menankaya |
| Cluster 10 | 32,...,45 | <i>S. soubrense</i> 'B' |
| Cluster 11 | 46,...,50 | <i>S. sanctipauli</i> 'Djodji' |
| Cluster 12 | 54,...,56,63,64 | <i>S. sanctipauli</i> |
| Cluster 13 | 57,59,...,62 | <i>S. sanctipauli</i> |

¹Numbers refer to the sample numbers given in Table 3.1, if a specific name is not given for the *a priori* taxon, then the species is *S. soubrense*.

Samples not included in the consensus classification:
 {4,5,51,52,53,58,65,66}

Table 3.7

Effect of the B_k methods on the distortion measure $\Delta_{0.5}$ for each *a priori* defined taxon.

| <i>a priori</i> taxon | B(k=) | | | |
|---|--------|--------|--------|--------|
| | 1 | 2 | 3 | 4 |
| <i>S. soubrense</i> 'B' | 0.6983 | 0.6588 | 0.6019 | 0.5153 |
| <i>S. sanctipauli</i> | 0.7408 | 0.5233 | 0.4227 | 0.4008 |
| <i>S. soubrense</i> Menankaya/Konkoure | 0.5696 | 0.529 | 0.4563 | 0.4332 |

SECOND PRINCIPAL AXIS

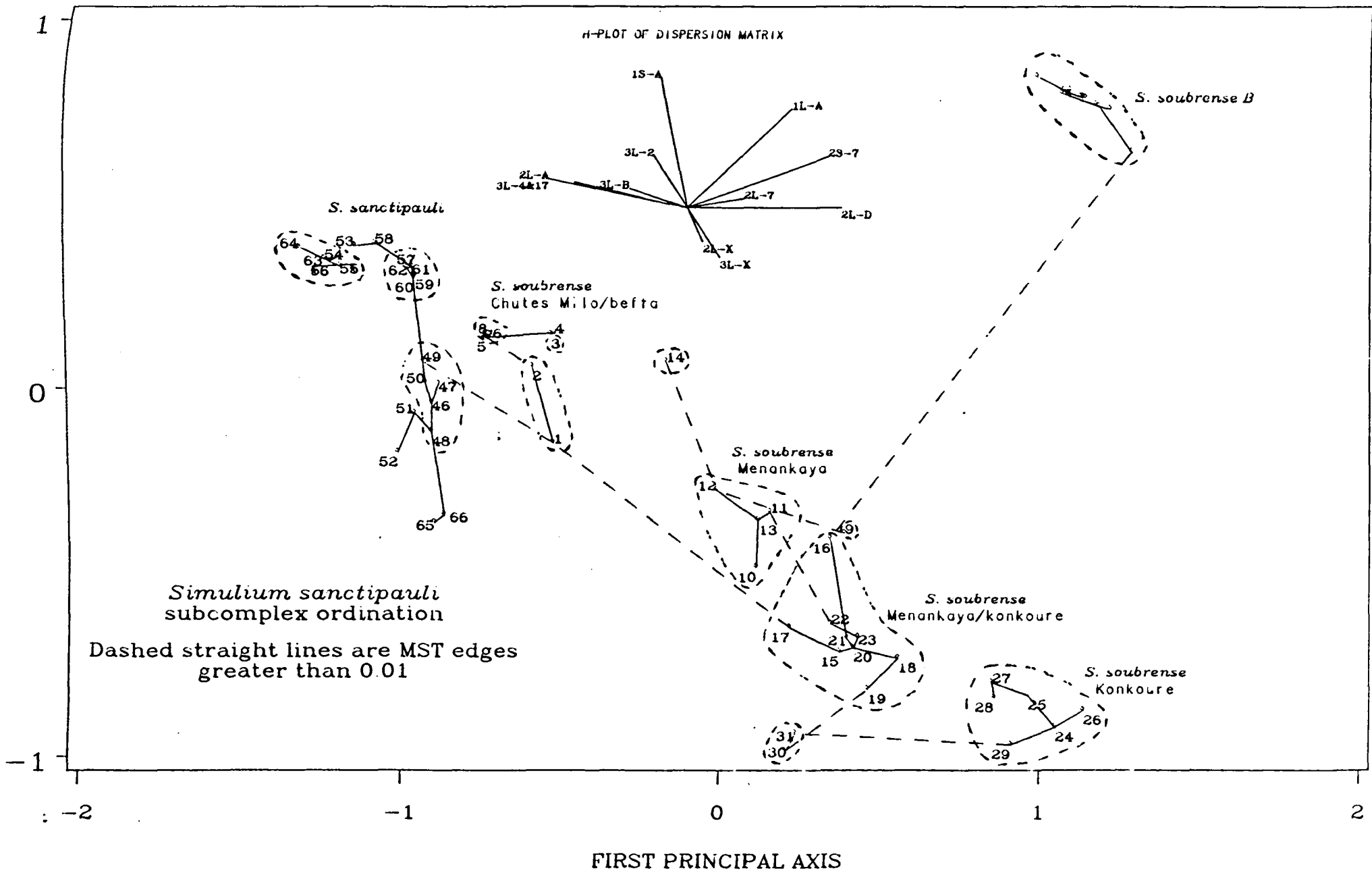


Figure 3.1

Figure 3.2
Hierarchical Cluster Analysis using Single Linkage

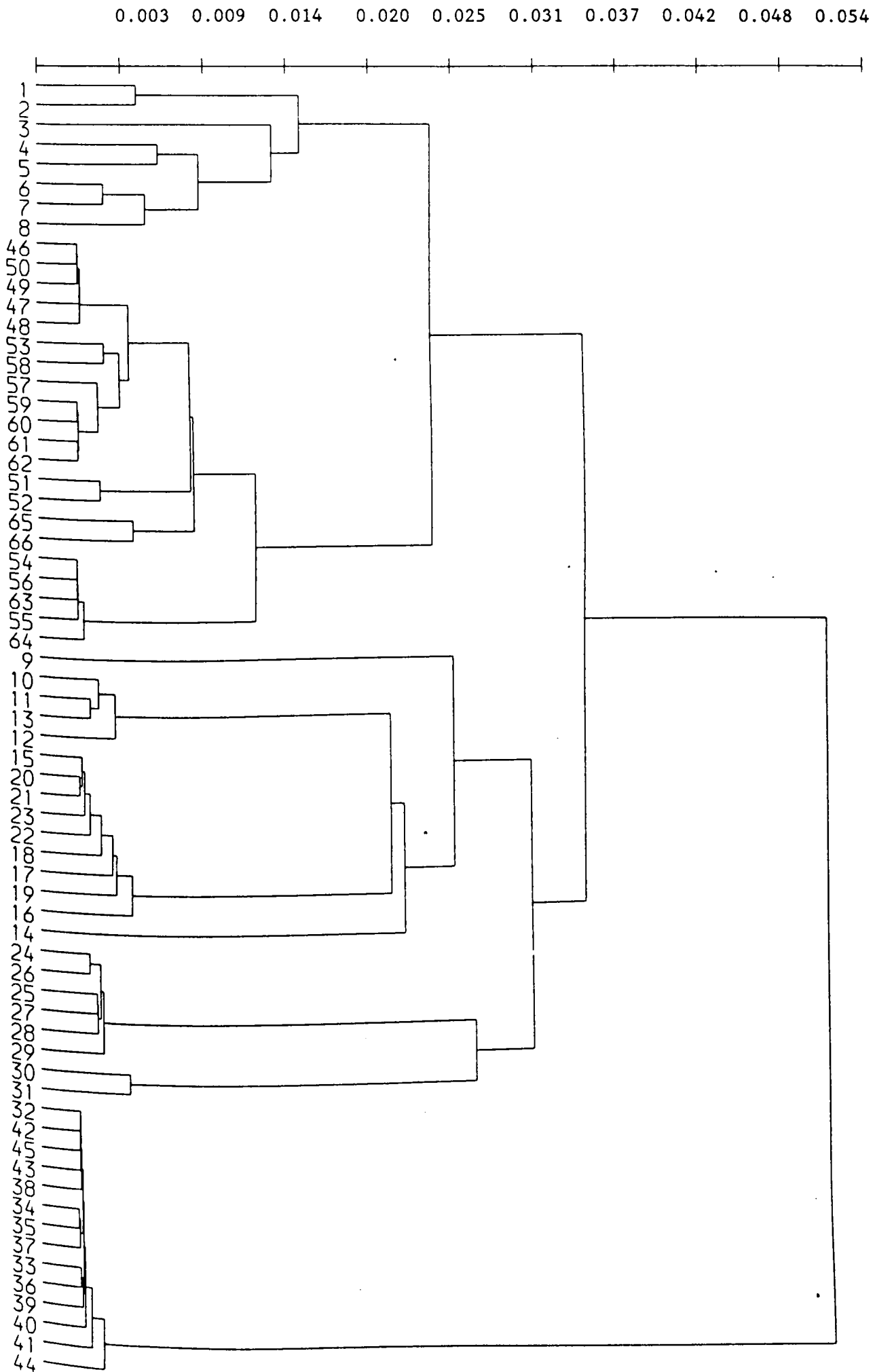


Figure 3.3
 Hierarchical Cluster Analysis using Complete Linkage

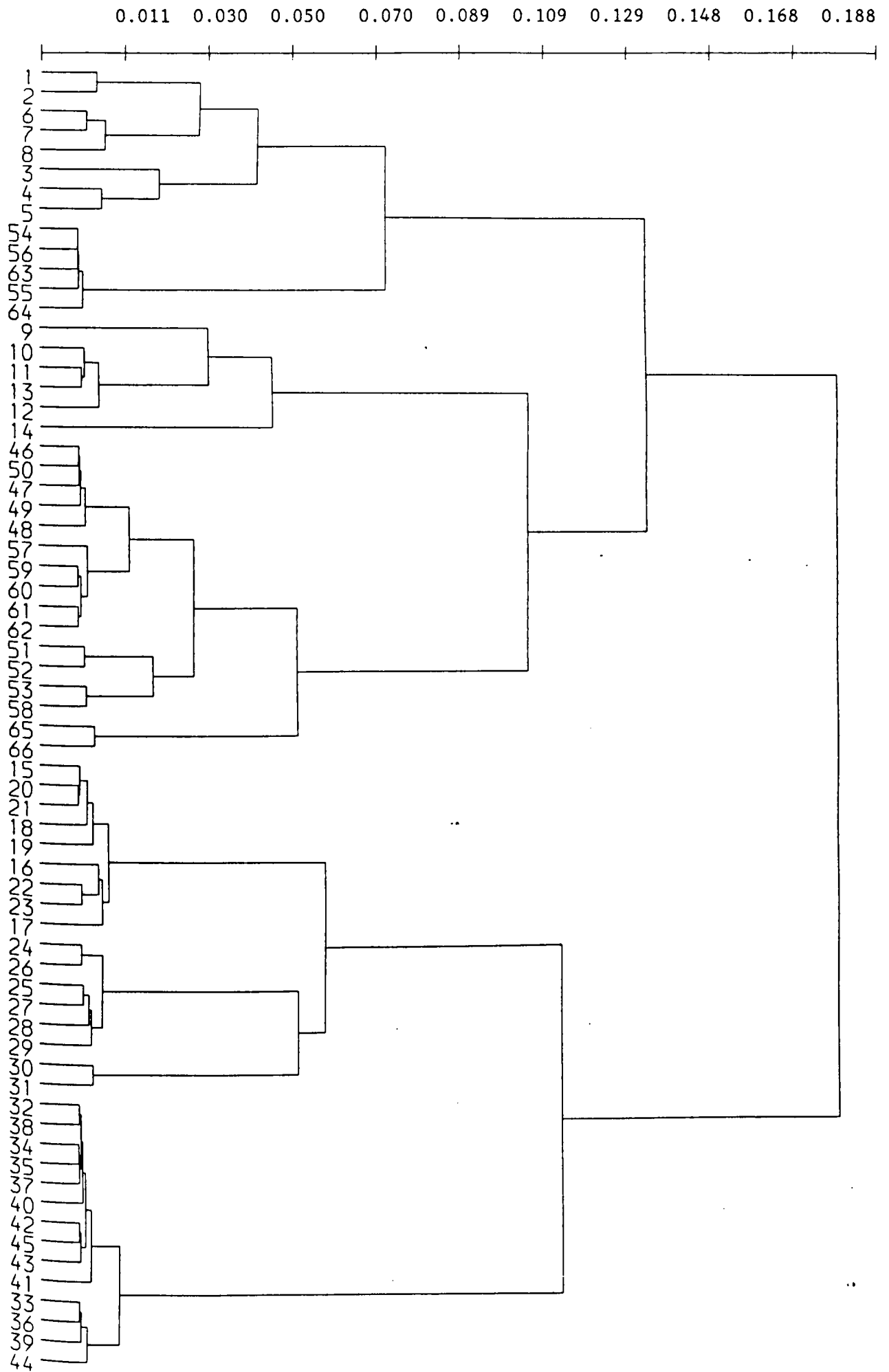


Figure 3.4
 Hierarchical Cluster Analysis using Average Linkage

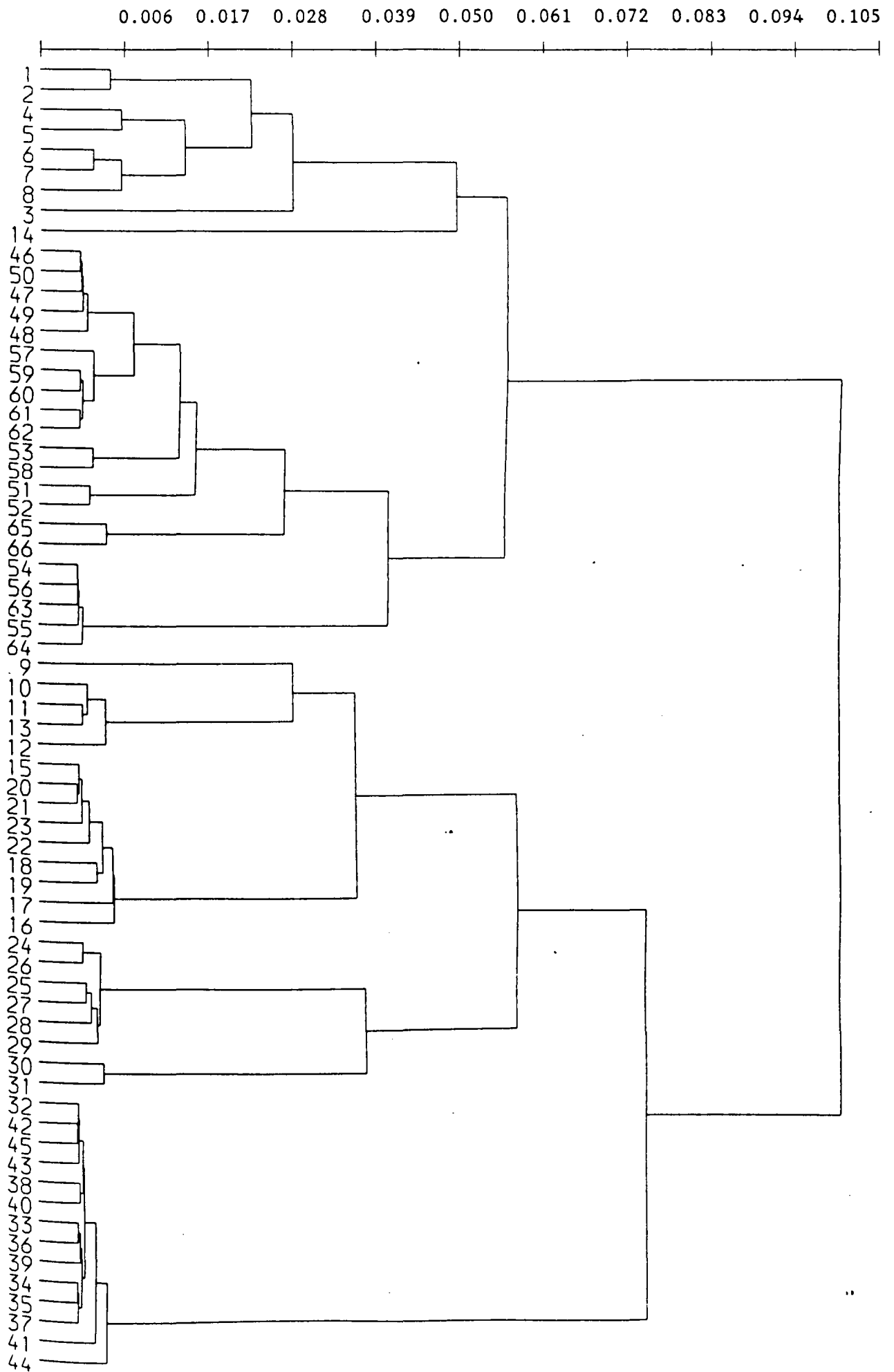
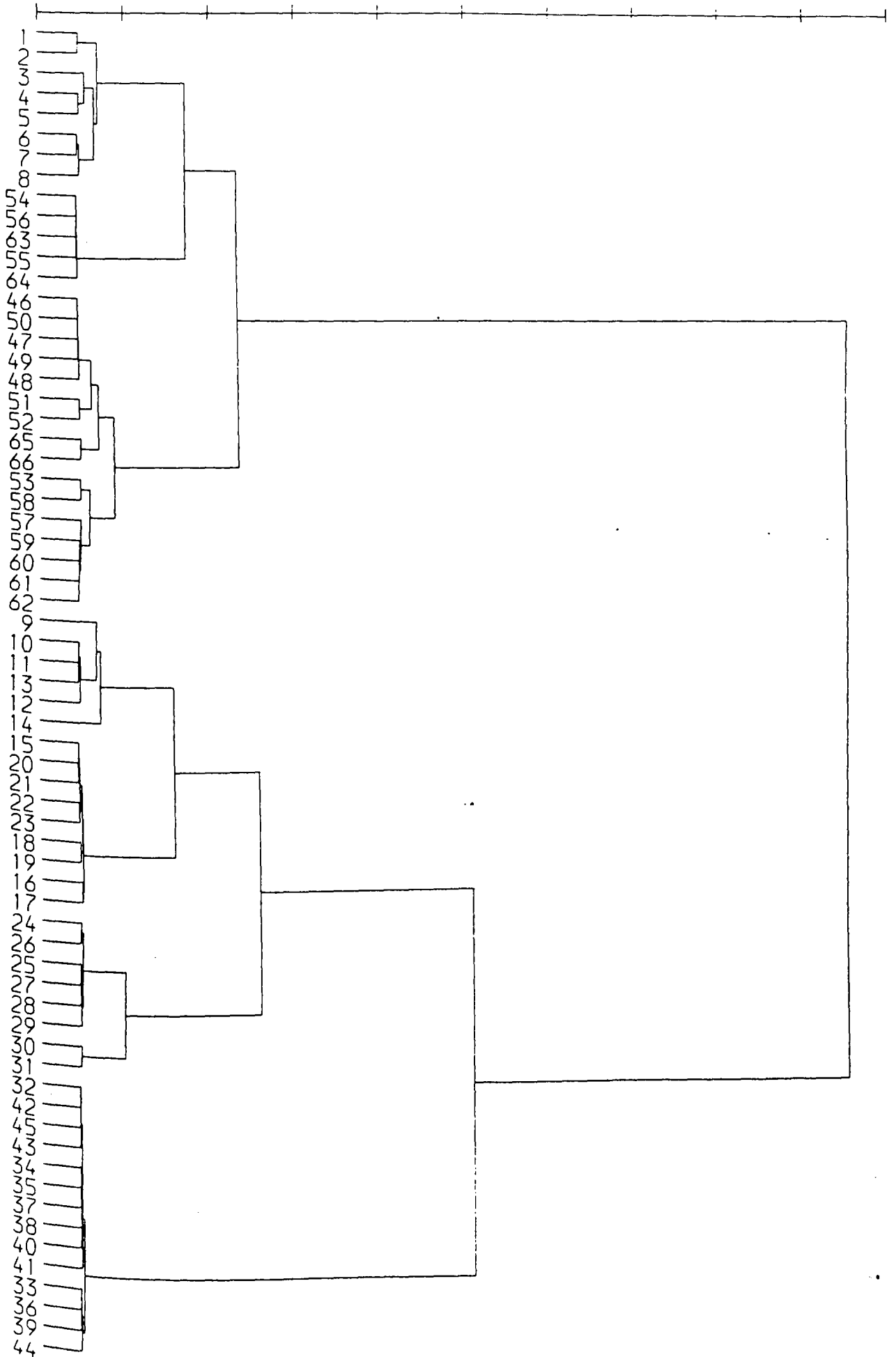


Figure 3.5
 Hierarchical Cluster Analysis using Ward's Method

0.112 0.317 0.522 0.727 0.932 1.137 1.342 1.547 1.752 1.957



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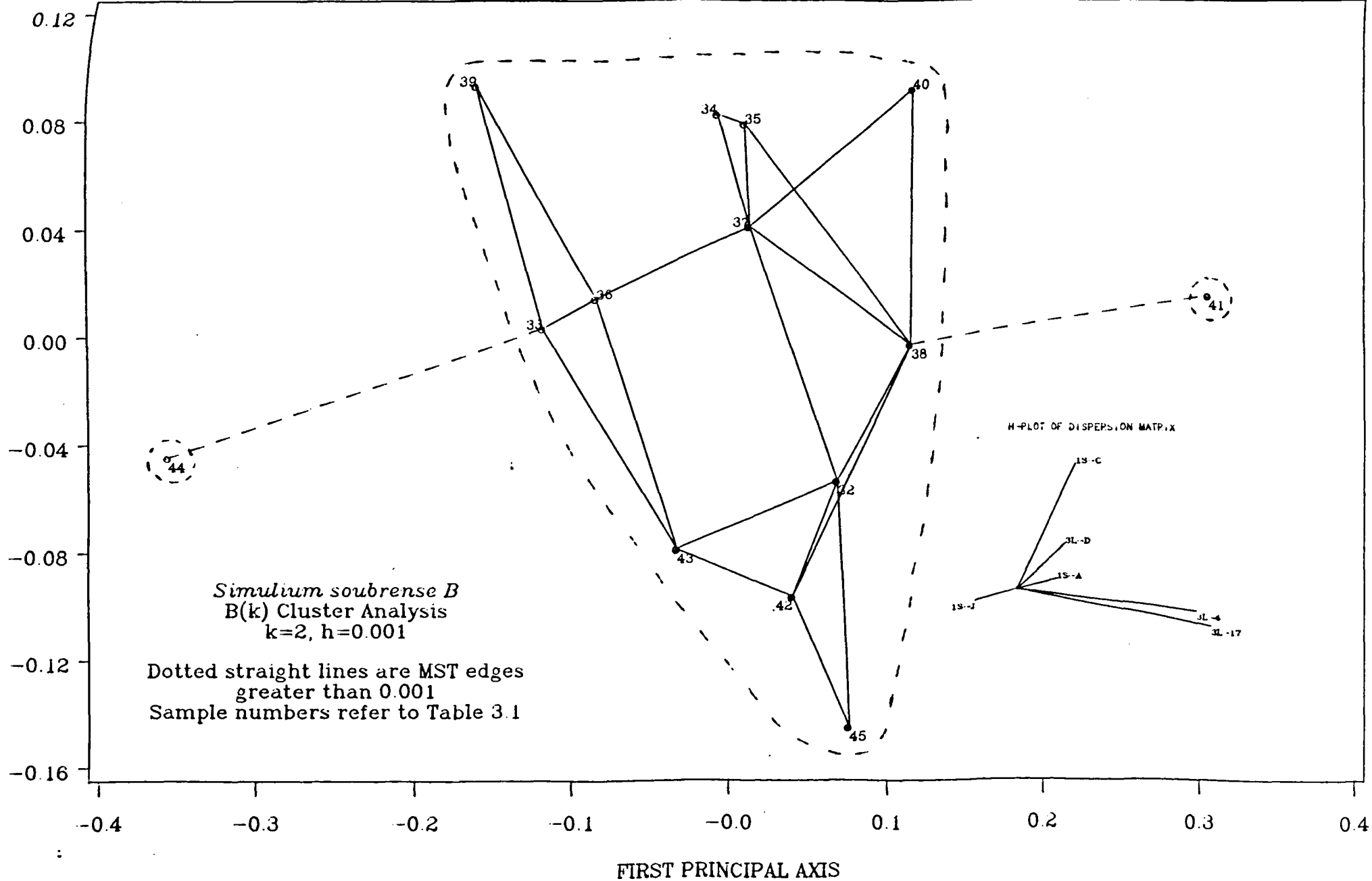


Figure 3.6

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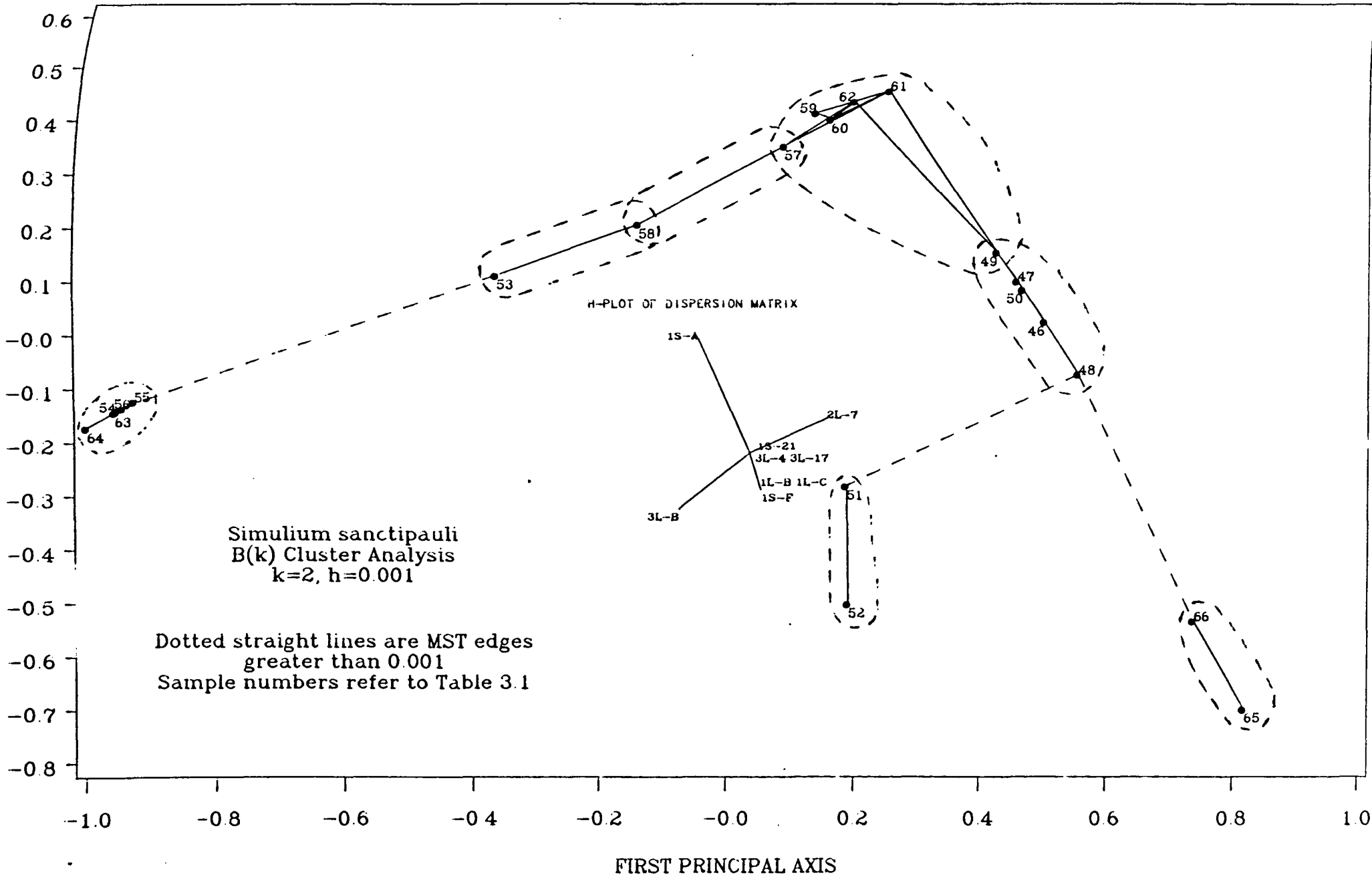


Figure 3.7

Simulium soubrense Menankaya/Konkoure
 B(k) Cluster Analysis
 $k=2, h=0.001$
 Dotted lines are MST edges
 greater than 0.001
 Sample numbers refer to Table 3.1

H-PLOT OF DISPERSION MATRIX

SECOND PRINCIPAL AXIS

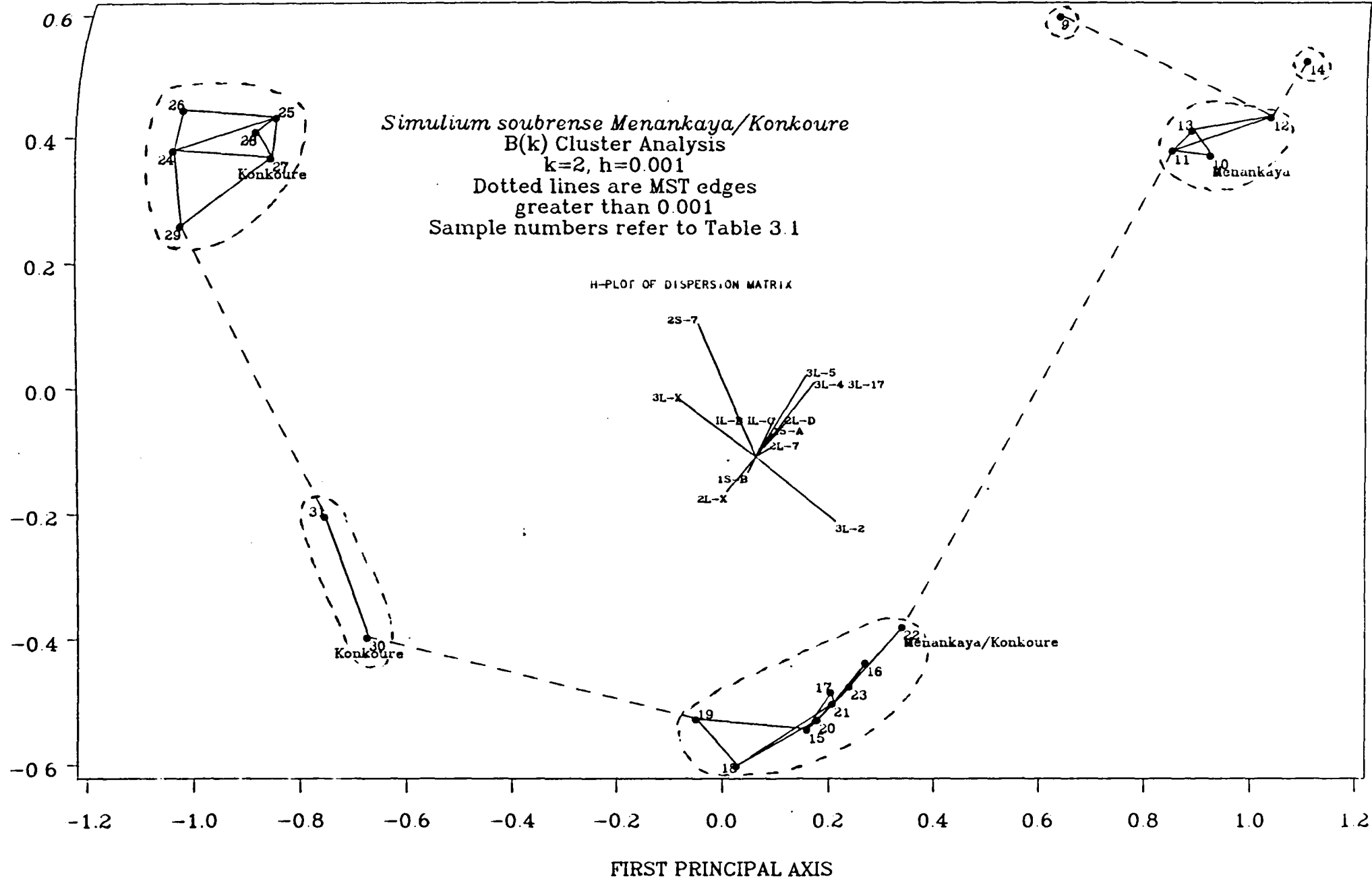


Figure 3.8

CHAPTER FOUR: ADULT FEMALE MORPHOMETRIC CHARACTERS

4.1 INTRODUCTION

The overall aim of the multivariate morphometric analysis of the *S. damnosum* complex is to provide a scheme for allocating an adult female *S. damnosum s.l.* to its correct cytospecies (Chapter one). The statistics used in this allocation scheme are developed in Chapters seven and eight, whilst Chapter nine describes the protocol to be followed for the actual allocation.

The purpose of the present chapter is to describe the morphometric characters which were used in the analyses following this chapter, and to describe how the characters were measured. The simple statistics used to describe each character were derived from the total data matrix following the procedures described in Chapter five.

4.2 MATERIALS AND METHODS

4.2.1 MATERIALS

Appendix one gives the full details of each sample used in the multivariate morphometric analyses presented in Chapters six, seven and eight. The initial sample size, and the final sample size following the methods given in Chapter five are shown.

4.2.2 METHODS

Appendix one describes the different sources of the samples of adult female *S. damnosum s.l.* available in this analysis. The females were either reared from pupae (which were collected attached to substrate at breeding sites and maintained at room temperature until

eclosion), with or without correlated cytotaxonomic identification (for cytotaxonomic methods see Chapter two), or caught at human bait. Each whole sample was then stored either in 95% propanol or in 70% ethanol, depending on the collector. The effect that storage in different media has on morphology was not evaluated in this analysis, but would certainly be worthy of future investigation.

Each adult female was examined in a culture dish in the same medium as it was preserved in, using a binocular microscope at 40X magnification. If the fly was missing for any of the characters described in the next section, it was discarded.

Four characters, thorax length, thorax width, clypeus width and head width, were measured at this stage, and the abdominal setal colouration recorded. In order to keep the fly in the correct plane, a plastic ring was glued to the base of the culture dish and the fly measured on this using an eyepiece graticule (Beech-Garwood *et al.* in the press). The head was removed for ease of measurement. The fly was then dissected, and the left antenna, left fore and mid legs, and the left wing mounted on a microscope slide in a small drop of Berlese mountant. After careful manipulation of the insect parts to ensure that all of the characters were clearly visible and not distorted, a cover slip was lowered at an angle onto the mounted parts. Only the left side of the insect was used to ensure that any variation due to systematic asymmetry of the flies was not detected.

The slide-mounted insect was then examined using a compound microscope at 32X, 100X, 250X and 400X magnification. The magnification for a particular character was used so that the character filled a large proportion of the eyepiece graticule. Measurements were recorded in graticule units on a standard record sheet, which contained

details of the sample from which the fly came, and a code number. The slide was marked with a code number and stored. The remainder of the fly was stored individually in the same medium as was used for the original sample.

At regular intervals during the data collection, data were typed into an IBM 3083 mainframe computer at the University of Liverpool Computer Laboratory. It was important that too much data were not input at any one time because the primary source of problems in the data derived from operator error (Chapter five). The data were edited in a temporary data set and carefully checked against the original record sheets. Once errors were corrected the temporary data set was appended to the main data set. The data were input in graticule units and a SAS data step program used to convert these units into microns, according to the calibration factor for each magnification, which was calculated from a standard stage micrometer slide.

4.2.3 ADULT FEMALE CHARACTER CHOICE

The choice of characters and the number of characters to be chosen is a problematic aspect of applied multivariate morphometric analysis (Blackith and Reyment 1971, p32). For technical reasons, the characters in this analysis were restricted to colour characters, counts and continuous linear characters. More sophisticated methods of direct shape assessment such as finite element scaling (Cherverud and Archie 1984) or Fourier methods (Rohlf and Archie 1984) could not be attempted. The problem of homology of characters does not usually arise in a closely related species complex, as it can in higher level taxa (Blackith and Reyment 1971). Thus a potentially enormous number of characters describing variation within and between members of the *S. damnosum* complex were available for measurement. Ideally both a

very large character set and very large sample sizes should be obtained in a multivariate morphometric analysis, so that no important character combinations are missed, and the estimation of parameters is accurate. In practise this was not possible, and a balance had to be struck between the number of characters and sample size.

Other practical considerations which were taken into account when choosing the initial characters were that the characters should:

1. be robust, i.e. not easily lost from flies,
2. be well defined, e.g. having a well defined landmark at each end of the body part;
3. be easily measured; i.e. a character was not included if it involved a large amount of time in dissection.

Besides these practical and statistical considerations, characters were also chosen which had been shown to have taxonomic value by earlier authors (Chapter one). Overall, a 28 eight character set was initially derived, which are individually described in section 4.3.

4.3 DESCRIPTION OF ADULT FEMALE CHARACTERS

An initial set of 28 characters was scored or measured for each individual fly. These characters were chosen according to the general principles described in section 4.2.3. The simple statistics reported for each character are those resulting from the total data matrix following the screening methods described in Chapter five. The relative taxonomic importance of these characters is assessed in Chapters six to nine. The characters are illustrated in Figures 4.1 to 4.7.

THORAX MEASUREMENTS

Two thorax characters were measured, length and width. It would have been preferable to have included a measure of thorax depth, so that the whole thorax volume was measured, but clearly defined landmarks were not available to make such a measurement reliable. In the past, thorax length has been used as a measure of overall size (Garms 1978), however it is unlikely that a single character can be wholly representative of size, and the thorax measurements were chosen in this analysis as characters in themselves.

4.3.1 THORAX LENGTH (V3)

The length of the thorax was measured from the posterior of the anterior thoracic spiracle to the middle of the posterior thoracic spiracle (Figure 4.1). This measure differs from the usual thorax length measurement (Garms 1978) which is the length from the anterior margin of the thorax to the posterior margin of the scutellum. This new measure was used because of the well-defined landmarks, which increases the accuracy of the measurement, although this new measure is smaller than the older measure, which in Garms (1978) ranged from 800 μ m to 1300 μ m. Thorax length measured in the old way is known to be a useful taxonomic character (Garms 1977), as a measure of size.

| | |
|--|----------------|
| Mean | 635.87 μ m |
| Standard error of the mean | 1.882 |
| Standard deviation | 53.243 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.207 |

4.3.2 THORAX WIDTH (V4)

The width of the thorax was measured across the greatest width of the scutum (Figure 4.2) This character has not been used in the morphometrics of the *S. damnosum* complex, and was chosen for its clearly defined landmarks.

| | |
|--|----------------|
| Mean | 882.08 μ m |
| Standard error of the mean | 2.3011 |
| Standard deviation | 65.0856 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.069 |

Figure 4.1

Thorax Length

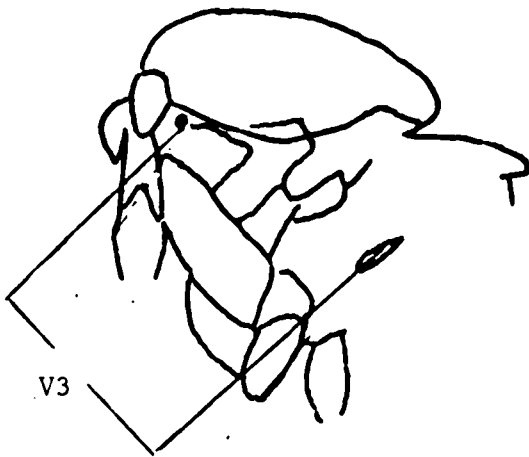
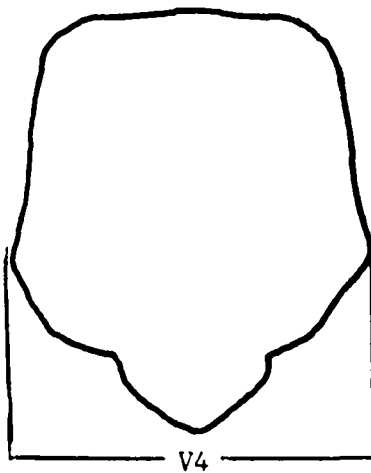


Figure 4.2

Thorax Width



HEAD MEASUREMENTS

Two head characters were measured, clypeus width and head width. Only these two head characters were measured because other characters did not have clearly defined landmarks. The number of maxillary teeth was not included in the analysis because it was an awkward character to score, and because Townson and Meredith (1979) and Garms (1978) both showed that it was not useful in distinguishing *S. yahense* form *S. squamosum*, as had been found by Quillévére and Sechan (1978). The palpal segments were not measured because they were difficult to dissect and were not linear.

4.3.3 CLYPEUS WIDTH (V5)

The width of the clypeus was measured between the extreme lateral points, just ventral to the antennae (Figure 4.3). The large proportion of samples for which the null hypothesis of normality was rejected was due to the low resolving power at 40X magnification for this small character. The sample distribution, whilst it was usually symmetrical was always stepped, due to rounding up or down of eyepiece graticule units. This character was therefore treated as being of questionable value prior to the analyses described in Chapters six to eight.

| | |
|--|-----------------------|
| Mean | 211.798 μm |
| Standard error of the mean | 0.5628 |
| Standard deviation | 15.919 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.862 |

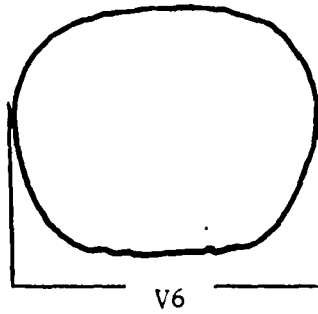
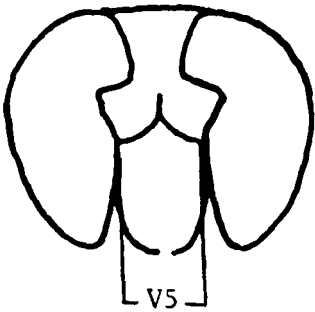
4.3.4 HEAD WIDTH (V6)

The width of the head was measured from the extreme lateral margins, viewed in posterior aspect (Figure 4.3). Anon (1976) measured head width in *S. damnosum s.l.* from Togo (Mean=874.26 μ m, s.e.=4.94), but found that it was not useful taxonomically.

| | |
|--|----------------|
| Mean | 806.46 μ m |
| Standard error of the mean | 1.779 |
| Standard deviation | 50.3101 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.0345 |

Figure 4.3

Clypeus and Head Width



COLOUR CHARACTERS

Dang and Peterson (1980) and Peterson and Dang (1981) presented a key to the species of *S. damnosum* s.l. based largely on colour characters. In general, colour characters were not included in this analysis because of the greater subjectivity involved with scoring such characters, and because the statistical procedures used in the analysis are more appropriate for analysing continuous linear characters. However two colour characters were included despite these problems because of their proven taxonomic importance in previous work.

4.3.5 WING TUFT COLOUR (V7)

The colour of the tuft of hairs at the base of the radial vein of the wing was recorded. The standard five state method was used for scoring the wing tuft colouration categories (Kurtak *et al.* 1981). The character states were as follows:

1. All pale hairs
2. Between one and five dark hairs (inclusive)
3. Mixed pale and dark hairs
4. Between one and five pale hairs (inclusive)
5. All dark hairs

The character was scored at high magnification using a compound microscope for greater accuracy.

Wing tuft colour has been used extensively in *S. damnosum* complex morphological studies in the past (Chapter one). In general it has been found that darker species are found in the forest while lighter species are found in the savanna. Some species (e.g. *S. squamosum*,

Kurtak *et al* 1981, *S. soubrense* 'Beffa', Meredith *et al.* 1983) are known to show all five character states.

4.3.6 ABDOMINAL SETAL COLOUR (V8)

The colour of the setae on the ninth abdominal tergite was recorded (Garms and Zillman 1984) as a two state character: character state two was defined as all hairs towards the middle of the tergite being long thick and dark, character state one was defined as anything other than this unless it was not possible to score the variable.

Previous work has shown this character to be 99.7% diagnostic for *S. yahense* (Garms and Zillman 1984), although more recent work has shown that it is 91% diagnostic (Thomson *et al.* 1987).

ANTENNAL MEASUREMENTS

A number of antennal measurements were made, reflecting the importance of this character in previous studies on the taxonomy of the *S. damnosum* complex (see Chapter one). Two measurements included more than one segment, measures V9 and V10. Neither of these included the third antennal segment as this was sometimes lost in dissection. V9 was recorded at a higher magnification than V10 which was more comprehensive of the whole antenna. The usual measure of antennal length (e.g. Garms 1978) is from the base to the tip.

In addition, five antennal segments were measured at 400X magnification. These characters were chosen because previous work has shown that the relative compaction of the antennal segments is taxonomically important (Quillévére *et al.* 1977, Garms 1978).

4.3.7 ANTENNAL LENGTH 1 (V9)

The antenna was measured from the base of the sixth to the tip of the eleventh antennal segment (Figure 4.4).

| | |
|--|-----------------------|
| Mean | 287.233 μm |
| Standard error of the mean | 1.111 |
| Standard deviation | 34.4266 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.0 |

4.3.8 ANTENNAL LENGTH 2 (V10)

The antenna was measured from the base of the fourth antennal segment to the tip of the eleventh antennal segment (Figure 4.4)

| | |
|--|-----------------------|
| Mean | 428.188 μm |
| Standard error of the mean | 1.6418 |
| Standard deviation | 46.438 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.0345 |

4.3.9 ANTENNAL SEGMENT 4 (V11)

The fourth antennal segment was measured at 400X magnification from the proximal to the distal margin (Figure 4.4).

| | |
|--|----------------------|
| Mean | 39.894 μm |
| Standard error of the mean | 0.199 |
| Standard deviation | 5.628 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.5172 |

4.3.10 ANTENNAL SEGMENT 5 (V12)

The fifth antennal segment was measured at 400X magnification from the proximal to the distal margin (Figure 4.4).

| | |
|--|---------------------|
| Mean | 39.26 μm |
| Standard error of the mean | 0.1935 |
| Standard deviation | 5.472 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.655 |

4.3.11 ANTENNAL SEGMENT 6 (V13)

The sixth antennal segment was measured at 400X magnification from the proximal to the distal margin (Figure 4.4).

| | |
|--|----------------------|
| Mean | 42.055 μm |
| Standard error of the mean | 0.2129 |
| Standard deviation | 6.0216 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.5172 |

4.3.12 ANTENNAL SEGMENT 7 (V14)

The seventh antennal segment was measured at 400X magnification from the proximal to the distal margin (Figure 4.4).

| | |
|--|----------------------|
| Mean | 41.387 μm |
| Standard error of the mean | 0.1989 |
| Standard deviation | 5.625 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.6897 |

4.3.13 ANTENNAL SEGMENT 8 (V15)

The eighth antennal segment was measured at 400X magnification from the proximal to the distal margin (Figure 4.4).

| | |
|--|----------------------|
| Mean | 40.785 μm |
| Standard error of the mean | 0.1905 |
| Standard deviation | 5.3873 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.7586 |

WING CHARACTERS

In all, six wing characters were measured or scored: wing tuft colouration and the ones presented in this section, of which four were continuous measurements and one a count.

4.3.14 WING LENGTH 1 (V16)

The length of the wing was measured from the humeral cross vein to the fusion of the subcostal and costal veins (Figure 4.5). This is not the whole wing length but only the length of the cell. This character has not been used in previous studies of the *S. damnosum* complex.

| | |
|--|-----------------------|
| Mean | 737.791 μm |
| Standard error of the mean | 2.0485 |
| Standard deviation | 57.939 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.0345 |

4.3.15 WING LENGTH 2 (V17)

The wing was measured from the point of fusion of the humeral cross vein and the costa to the radius media crossvein (Figure 4.5).

| | |
|--|-----------------------|
| Mean | 450.943 μm |
| Standard error of the mean | 1.222 |
| Standard deviation | 34.567 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.069 |

4.3.16 WING WIDTH 1 (V18)

The width of the wing was measured from the point of fusion of the two medial veins to the meeting point of the second cubital vein and the wing margin (Figure 4.5).

| | |
|--|------------------------|
| Mean | 1010.969 μm |
| Standard error of the mean | 2.489 |
| Standard deviation | 70.402 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.2414 |

4.3.17 WING WIDTH 2 (V19)

The width of the wing was measured from the meeting point of the second cubital vein and the wing margin to the end of the radial sector (Figure 4.5).

| | |
|--|-------------------------|
| Mean | 1428.8996 μm |
| Standard error of the mean | 3.5159 |
| Standard deviation | 99.443 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.1034 |

4.3.18 WING LENGTH 3 (V20)

The length of the wing was measured from the end of the radial sector to the radius media crossvein (Figure 4.5).

| | |
|--|------------------------|
| Mean | 1487.153 μm |
| Standard error of the mean | 3.791 |
| Standard deviation | 107.217 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.06897 |

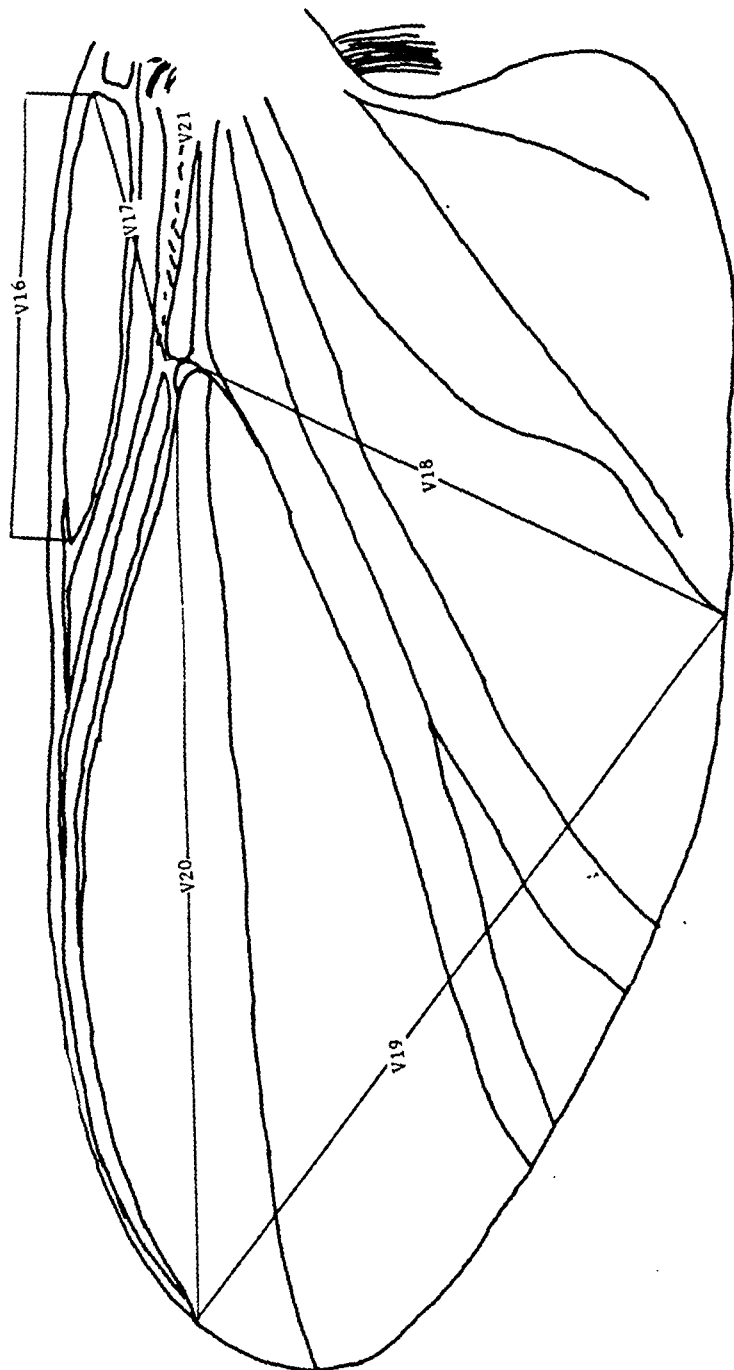
4.3.19 NUMBER OF HAIRS ON THE RADIAL VEIN (V21)

The number of hairs on the radial vein up to the radius media cross vein were counted (Figure 4.5). This character was previously used by Quillévéré and Sechan (1978) as a taxonomic character in the *S. damnosum* complex, but was subsequently shown to be of little taxonomic value by Garms (1978) and Townson and Meredith (1979).

| | |
|--|--------|
| Mean | 15.274 |
| Standard error of the mean | 0.1192 |
| Standard deviation | 3.371 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.3103 |

Figure 4.5

ing Characters



LEG CHARACTERS

Nine leg characters were measured or scored, eight continuous measurements and one a count. Leg shape has previously been shown to be variable within *S. damnosum s.l.* (Lewis and Duke 1966). Anon (1976 and Soponis and Peterson 1976) also measured a number of leg segments and found them to be taxonomically useful.

4.3.20 FEMUR LENGTH 1 (V22)

The length of the femur of the left fore leg was measured in anterior aspect from the strongly sclerotized ventral articulation to the extreme dorsal end (Figure 4.6). Anon (1976) measured a similar character (Mean=616.87 μ m, s.e.=3.5) and considered it to be taxonomically useful.

| | |
|--|-----------------|
| Mean | 633.352 μ m |
| Standard error of the mean | 1.7022 |
| Standard deviation | 48.144 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.1034 |

4.3.21 TIBIA LENGTH 1 (V23)

The length of the tibia of the left fore leg was measured in anterior aspect from the extreme dorsal end of the proximal end to the end of the tibia (Figure 4.6). Anon (1976) measured this character (Mean=695.87 μ m, s.e.=3.91) but did not consider it to be taxonomically useful.

| | |
|--|-----------------|
| Mean | 695.971 μ m |
| Standard error of the mean | 1.7783 |
| Standard deviation | 50.2925 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.1724 |

4.3.22 BASITARSUS LENGTH 1 (V24)

The length of the basitarsus of the left fore leg was measured in anterior aspect from the point of articulation with the tibia to the distal end (Figure 4.6). Anon (1976) measured this character (Mean=437.75 μ m, s.e.=1.03) and considered it to be taxonomically useful.

| | |
|--|----------------|
| Mean | 438.79 μ m |
| Standard error of the mean | 1.202 |
| Standard deviation | 39.999 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.1035 |

4.3.23 TARSAL SEGMENT 2 (V25)

The second tarsal segment of the left fore leg was measured from its proximal to its distal end (Figure 4.6). Anon (1976) measured this character (Mean=161.5 μ m, s.e.=1.03) but did not consider it to be taxonomically useful.

| | |
|--|-----------------|
| Mean | 165.769 μ m |
| Standard error of the mean | 0.425 |
| Standard deviation | 12.02 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.0178 |

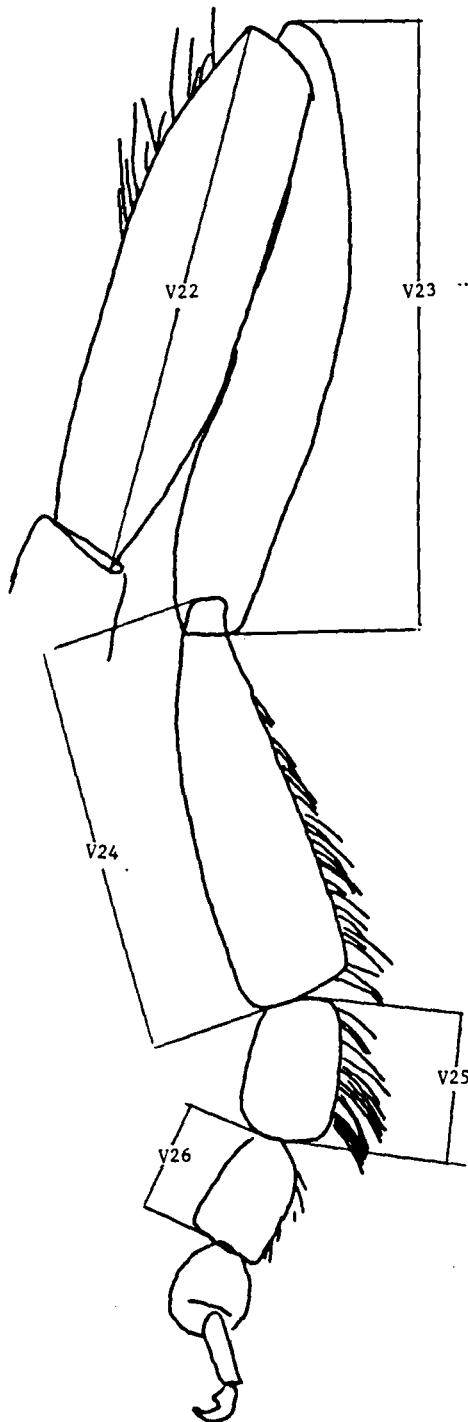
4.3.24 TARSAL SEGMENT 3 (V26)

The third tarsal segment of the left fore leg was measured in anterior aspect from its proximal end to its distal end (Figure 4.6). Anon (1976) measured this character (Mean=120.1 μ m, s.e.=1.03) but did not consider it to be useful.

| | |
|--|-----------------|
| Mean | 124.605 μ m |
| Standard error of the mean | 0.3526 |
| Standard deviation | 9.9744 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.069 |

Figure 4.6

Fore-leg Characters



4.3.25 FEMUR LENGTH 2 (V27)

The length of the femur of the left mid-leg was measured in posterior aspect from the ventral articulation with the coxa to the dorsal distal end (Figure 4.7). Anon measured this character (Mean=628.82 μ m, s.e.=3.6) but did not consider it to be taxonomically useful.

| | |
|--|-----------------|
| Mean | 663.247 μ m |
| Standard error of the mean | 1.7036 |
| Standard deviation | 48.185 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.138 |

4.3.26 TIBIA LENGTH 2 (V28)

The length of the tibia of the left mid-leg was measured in posterior aspect from the dorsal/proximal articulation with the femur to the distal end (Figure 4.7). Anon (1976) measured this character (Mean=620.99 μ m, s.e.=3.6) but did not consider it useful.

| | |
|--|-----------------|
| Mean | 623.229 μ m |
| Standard error of the mean | 1.5927 |
| Standard deviation | 45.047 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.1035 |

4.3.27 BASITARSUS LENGTH 2 (V29)

The basitarsus of the left mid-leg was measured in posterior aspect from the point of articulation with the tibia to the distal end (Figure 4.7). Anon (1976) measured this character (Mean=316.72 μ m, s.e.=2.01) and considered it to be taxonomically useful.

| | |
|--|-----------------|
| Mean | 328.755 μ m |
| Standard error of the mean | 0.996 |
| Standard deviation | 28.1706 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.0345 |

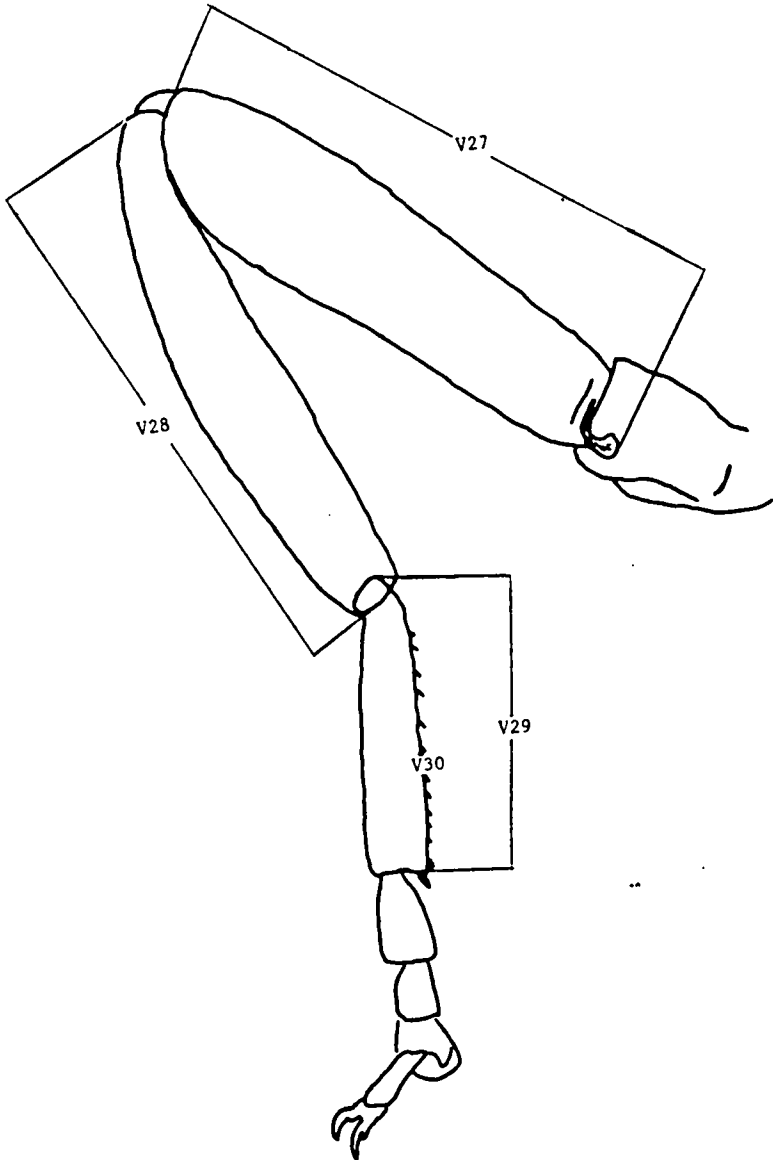
4.3.28 BASITARSAL SPINE NUMBER (V30)

The number of spines on the dorsal margin of the basitarsus of the left mid-leg were counted (Figure 4.7). The sample size for this character was smaller (N=777) than all the other characters in this section (N=800) because this character was rejected from the analysis at an early stage (Chapter 5) due to a large number of missing values and poor discriminatory power.

| | |
|--|--------|
| Mean | 24.51 |
| Standard error of the mean | 0.1122 |
| Standard deviation | 3.128 |
| Proportion of samples for which the null hypothesis of normality was rejected, $\alpha=0.05$ | 0.1724 |

Figure 4.7

Mid-leg Characters



4.4 DISCUSSION

The character set chosen for this analysis comprised 28 characters, of which 15 were similar or identical to characters previously considered taxonomically useful in the *S. damnosum* complex. The number of characters measured or scored could in principle have been extended indefinitely, but this would have severely restricted the range of variation which could have been sampled.

The time taken to measure or score all 28 characters was roughly 20 minutes. The lengthiest measurements to take were those taken on the dissecting microscope because it was more difficult to ensure that the fly was positioned in the correct plane, while those characters measured on the compound microscope were slide-mounted and so in the correct plane. Dissection of the fly was also time consuming, and was made more or less difficult according to the preservative in which the fly was stored. Ethanol preserved material was less brittle than Propanol preserved material. Propanol was preferred because the brittleness of the fly made dissection quicker and easier.

CHAPTER FIVE: WITHIN SAMPLES VARIATION

5.1 INTRODUCTION

The major objective of the multivariate morphometric analysis of the *S. damnosum* complex is to develop a technique for the identification of adult female *S. damnosum* s.l. to their correct cytospecies. To develop the appropriate statistics for this method it is necessary to obtain sufficient samples of each cytospecies which together comprehensively cover the range of variation within each cytospecies.

The purpose of this chapter is to describe the statistical methods used to ensure that the variation observed within each sample was not unduly influenced either by mistakes within the data set or by rogue individuals.

There were two basic requirements which the available samples needed to meet for them to be included in the analyses presented in Chapters six to nine: the chromosomal identity of the sample was known, and the individuals within the samples were not atypical of the range of variation found within the cytospecies.

Examining Appendix one shows that the strength of evidence supporting the chromosomal identity of samples ranged from the moderately strong to the relatively weak. Thus some samples were known from correlated larval cytotaxonomy to be pure for one species (e.g sample V1=1, *S. squamosum*), or to be a mixture of species (e.g sample V1=5, *S. squamosum* and *S. sanctipauli*); other samples did not have correlated larval identifications but were believed, from other evidence to be a certain cytospecies (e.g. sample V1=15, *S. squamosum*) or

cytospecies mixture (e.g. sample V1=10, *S. damnosum* s.s., *S. squamosum*/*S. yahense* from correlated DNA probes identifications).

The pretreatment that a sample underwent depended on the *a priori* information available as to its species composition and purity. In the case of pure samples, the problem of pretreatment was restricted to the identification of outliers and occasional contaminants, whilst in the case of mixed samples the problem was one of separating a mixture of distributions followed by outlier detection in the separated parts.

5.2 MISSING VALUES

Not all individuals could be measured or scored for every character described in Chapter four, either because of damage to or loss of the character during dissection, or because of distortion during the slide-mounting stage. Some characters were missing more often than others (Appendix one). This created two data-analytic problems, firstly what to do with a character for which missing values were relatively common, and secondly what to do with individual flies with one or more missing values.

Characters which could not be scored for a relatively large number of individuals were tested in the course of the data-collection phase of the analysis (i.e. prior to the separation of mixtures or the detection of outliers) to see if they had any discriminatory power, using one-way ANOVA and discriminant analysis. If the character proved to be taxonomically unimportant relative to the number of missing values, it was discarded from subsequent analyses and individuals missing for only that character but complete for the rest were brought into the main data set. This procedure resulted in the basitarsal spine number being rejected due to the combination of a large number of missing values and poor discriminatory value. By contrast, wing tuft colouration was kept because of its good discriminatory value in spite of a relatively large number of missing values.

Any fly missing for any of the remaining 27 characters in the analysis was rejected from the main data set and not used again.

5.3 UNIVARIATE OUTLIERS

Once a sample was cleared of flies with missing values, each character was screened for univariate outliers. Barnett and Lewis (1984) define an outlier as an observation (or subset of observations) which appears to be inconsistent with the remainder of the data set. Outlying individuals can arise in a number of ways (Barnett and Lewis 1984) including:

1. operator error (i.e. typing error, recording error, measurement error);
2. as outliers in relation to a specific model (e.g. normal distribution);
3. as contaminants from other distributions.

The underlying model assumed for the characters in this analysis was the univariate normal distribution (unless the character departed from this model in a large proportion of samples, see Chapter four). Outliers were assessed using the following methods.

5.3.1 INFORMAL GRAPHICAL METHODS

There are many informal graphical methods for examining a sample distribution (Tukey 1977, Seber 1984) Amongst the simplest is the 'stem-and-leaf' plot (Mosteller and Tukey 1977), which is an informal histogram, where the 'stem' is equivalent to the midpoints of a standard histogram and the 'leaves' represent the number and value taken by individuals within the class interval defined by the midpoints. This plot was used to assess the shape of the sample distribution, in particular its symmetry and the length of the tails, aspects of the distribution which have most influence on the usual estimates of parameters (Barnett and Lewis 1984). 'Box-and-whisker' plots (Tukey 1977) were also used in this analysis to assess the same

sample properties (Figure 5.1). This is a graphical representation of the data which identifies the mean, median, inter-quartile range, and outlying individuals. Finally, the ordered data for each sample were plotted against the quantiles of the standard normal distribution $N(0,1)$, giving a quantile-quantile (Q-Q) plot (Gnanadesikan 1977). A straight line results if the data are sampled from a normal distribution. This method was a useful informal method of assessing distributional assumptions and identifying outlying individuals.

These methods were all output as part of the procedure SAS PROC UNIVARIATE. Figure 5.1 is a sample of the output from such an analysis showing the three plots used. The individuals identified as outliers in this example were from sample V1=17 (Appendix one), the character was thorax length. In practise these methods resulted in voluminous printed output, so to be practicable most analyses were performed interactively, using the Display Manager System within SAS.

5.3.2 FORMAL STATISTICAL TESTS

Whilst the emphasis at this stage of the analysis was on informal data-analytic methods, three statistical tests were performed. Graphical methods are very useful but it is quite easy to miss important features of a graphical representation of the data, especially if the analysis is largely interactive. The following tests were used:

1. Shapiro-Wilk W test (Seber 1984,p452) for univariate normality; If the null hypothesis of normality was rejected, the graphical plots described in section 5.3.1 were examined, as departure from normality can result from the following reasons:

- a). genuine non-normality (i.e. the sample is drawn from another distribution, such as the log-normal distribution)

- b). inadequate measurement of a normally distributed character such as clypeus width (Chapter four), identified by approximate symmetry, ordinary tails but a stepped distribution
- c). from the presence of outliers.

2. Single upper or lower outlier test (Barnett and Lewis 1984, p167 Table VIII)

3. Upper and lower outlier pair test (Barnett and Lewis 1984, p171, Table XIIa).

These two outlier test were chosen for their power and relative simplicity to calculate.

When an individual was identified as an outlier using the formal and informal methods, the data were checked for typing errors. This was a common source of outliers in the data, and was also the simplest to correct. If, however, the record sheet and the data on the computer agreed, then the original measurement was checked for recording error, and the individual re-measured for that character. If a fly appeared to be genuinely discordant then it was removed from the data set and the whole procedure repeated until all such individuals were identified. Appendix one gives the number of flies rejected from each sample as a result of either univariate or multivariate outliers.

5.4 MULTIVARIATE OUTLIERS

The detection of univariate outliers described in the previous section is a relatively straightforward task, involving a simple ordering of the data. However outlier detection in multivariate data is an extremely complex procedure for the following reasons:

1. there is no unique ordering of multivariate data;
2. graphical representation of high dimensional data is not easy;
3. outliers in multivariate data can distort the association between characters (correlation) as well as the estimates of location and spread affected by univariate outliers. (see Seber 1984, Barnett and Lewis 1984, Gnanadesikan 1977, Gnanadesikan and Kettenring 1972 for details). In addition bias is introduced due to sampling error in the estimate of the parameters of the joint multivariate distribution for a small sample size and moderately large dimensionality (Seber 1984). Because of these problems a number of different methods were used to identify multivariate outliers in an informal, interactive approach.

5.4.1 PRINCIPAL COMPONENTS ANALYSIS

One way of imposing order on multivariate data is to find linear combinations of the original characters (Seber 1984) and to examine the ordering of individuals along these new variables. A straightforward and well tried method for finding linear combinations is principal components analysis (Seber 1984, Gnanadesikan 1977). The first principal component is that linear combination of the original variables along which the variance is a maximum, subsequent principal components account for progressively less variance until r principal components have been extracted (where r is the rank of the dispersion (variance/covariance) matrix, equal to the number of characters, p ,

if the dispersion matrix is not singular). For data sampled from a multivariate normal population, the scatter of points along each component is expected to be univariate normally distributed (Mardia, Kent and Bibby 1977, p230), and so can be examined using the graphical and statistical methods described in section 5.3. Furthermore, the component on which an individual is an outlier gives information as to the nature of the distinctiveness of the fly (Gnanadesikan 1977). For example, if an outlier was found on the first principal component, usually a size vector in multivariate morphometric analysis (Blackith and Reyment 1971), then its distinctiveness was due to systematic differences in all characters simultaneously (i.e. size), whereas an outlier on one of the shape vectors could be due to the association between one or more characters (i.e. shape).

Principal components are extracted either from the sample dispersion (variance/covariance) matrix or from the correlation derived from it. Because the estimates of dispersion are very sensitive to outliers it has been suggested that a robust estimate of dispersion matrix (i.e an estimate which is insensitive to outlying individuals, Huber 1981) should always be calculated and the principal components extracted from this (Gnanadesikan 1977). This was attempted using a method for calculating robust principal components (Campbell 1980, using a GENSTAT macro written by Matthews 1984). This method has the effect of making discordant observations more obvious, because observations away from the bulk of the data are given reduced weight, and so influence the estimates of mean and dispersion to a smaller extent than usual. The method was very expensive computationally and could not be used in the interactive way that the usual principal components analysis computed by SAS PROC PRINCOMP could.

To summarise, each sample was subjected to a principal components analysis, the scatter of points along the principal components were examined for normality and the presence of outliers. Bivariate scatters of selected principal components were also examined, in particular the first two (accounting for most variance) and the last two (accounting for least variance). Characters which were known *a priori* not to be univariate normally distributed were not usually included in the principal components analysis. As a result of this method a list of multivariate outliers for a sample was obtained.

5.4.2 MAHALANOBIS' DISTANCE

The Mahalanobis' squared distance of each individual to the sample mean can be used to order the individuals in a sample. The distance: $(x_i - \bar{x})' S^{-1} (x_i - \bar{x})$ where x_i is the vector of observations for the i -th individual, \bar{x} is the mean vector of the sample, and S is the sample dispersion matrix, is distributed as a χ^2 variable with p (number of variables) degrees of freedom (Seber 1984). For p up to and including five, a discordancy test for a single multivariate normal outlier was used (Barnett and Lewis 1984, Table XXX). Higher dimensionalities than this could not be tested for discordant individuals in this way. Instead, an informal graphical method was used, which involved plotting the ordered Mahalanobis' squared distances against the quantiles of a Beta distribution with parameters $a = \frac{1}{2}p$ and $b = \frac{1}{2}(n-p-1)$, where n is the sample size and p is the number of characters. If the data were multivariate normal, the result was a straight line (Seber 1984, p153). Figure 5.2 is an example of such a plot, with the outlier clearly visible on the top Q-Q plot. Computationally this was performed using a routine written in SAS PROC MATRIX. A simpler Q-Q plot was also drawn by plotting the cube roots of the ordered distances

against the quantiles of a standard normal distribution (Campbell 1984). These plots were used interactively to assess the distributional properties of the sample and to identify discordant individuals within the sample, however the high dimensionality and small sample size often rendered the results of this method equivocal, and so it was usually used only on relatively small numbers of characters at a time.

As well as these plotting methods, Mardia's (1970) estimates of multivariate skewness and kurtosis were calculated in the same SAS PROC MATRIX routine, as these could be derived in earlier steps of the program to calculate the Mahalanobis' squared distances. These statistics have poor small sample properties (Seber 1984), but were used as general indicators of the adequacy of the data set. The multivariate kurtosis statistic was especially useful in the identification of multivariate outliers, to which it is particularly sensitive.

Using these techniques, a list of potential multivariate outliers was obtained for each sample. At each stage, the individual which was the most 'extreme' outlier was discarded, until all outliers were identified and excluded from the data set.

5.5 MIXED SAMPLES

A number of samples were known *a priori* to be mixtures of cytospecies (Appendix one), from correlated larval cytotaxonomy. Therefore, before the set of pure-sample outlier detection techniques described in previous sections could be applied it was necessary to separate mixed samples into their constituent species. Two approaches to this problem were taken: internal analysis and external analysis.

5.5.1 INTERNAL ANALYSIS

With internal analysis, only the sample data were used to separate the mixture into its constituent parts. A number of dimension reduction methods were used to view the data comprehensively, and so identify groupings within the data (Gordon 1981), including principal components analysis and cluster analysis methods (Gordon 1981, Chapter three). Using these methods, the sample was split into its component parts and the parts identified as a particular species with reference to the expected proportions from correlated cytotaxonomy, or from external analysis (section 5.5.2).

5.5.2 EXTERNAL ANALYSIS

With external analysis, the data used for separating the mixture came from outside of the sample data. For a small number of samples, independent DNA probes identifications were available (Post pers. comm.), whilst other samples were separated using linear discriminant functions calculated from other samples belonging to the species known to be contained in the mixed sample.

Once a mixture was separated using these techniques, each part was then examined as if it were a pure sample, for univariate and multivariate outliers using the methods described in sections 5.3 and 5.4.

5.6 SUMMARY OF METHODS

For samples which were known *a priori* to be pure, the initial sample was examined for missing values in the 27 characters for which missing values were to be rejected. These flies were discarded from the main data set. Univariate outlier tests were applied to each character in turn, typing and recording errors were recognised and remedied, any flies which were outlying were rejected. Multivariate outlier tests were then applied, either to the whole character set or to subsets of characters, and any multivariate outliers which were identified were rejected.

For samples which were known *a priori* to be mixed, the exact procedure used depended upon the proportion of each species in the mixture. Thus if one species was present only as a small proportion of the total sample, the methods applied to the sample were basically those applied to a pure sample with outliers and contaminants. Initially, missing values were identified and rejected, and typing or recording errors corrected. Then the internal (section 5.5.1) and external analyses (section 5.5.2) were applied as appropriate, any flies which did not fall clearly into any of the clusters revealed by these methods were rejected, and the initial sample divided into its component parts. The 'new' samples produced in this way were then treated in exactly the same way as if they were pure samples, with univariate and multivariate outliers being detected and rejected.

5.7 DISCUSSION

The methods described in this chapter aimed to identify problems within each sample of *S. damnosum s.l.* and to solve them either by correcting the data, or by removing rogue individuals. To some extent the choice of flies to be rejected involved some subjectivity, because some methods identified flies as outliers which were not recognised by other methods. However, this problem was of less importance than the fact that obviously influential flies were identified by all methods without equivocation, so that the minimum effect of applying the methods described in this chapter was to identify flies whose presence in the data set would have seriously affected the subsequent analyses.

Figure 5.1

Example printout from SAS PROC UNIVARIATE showing stem-and-leaf plot, boxplot and Q-Q plot, to identify an outlier.

| STEM LEAF | # | BOXPLOT |
|----------------|----|------------|
| 74 5 | 1 | |
| 72 333 | 3 | |
| 70 888 | 3 | |
| 68 33333336666 | 11 | +-----+ |
| 66 11111 | 5 | *--*--*--* |
| 64 68888 | 5 | +-----+ |
| 62 114 | 3 | |
| 60 9 | 1 | |
| 58 66 | 2 | |
| 56 | | |
| 54 9 | 1 | 0 |
| 52 | . | |
| 50 | | |
| 48 7 | 1 | * |

-----+-----+-----+-----+
 MULTIPLY STEM.LEAF BY 10**+01

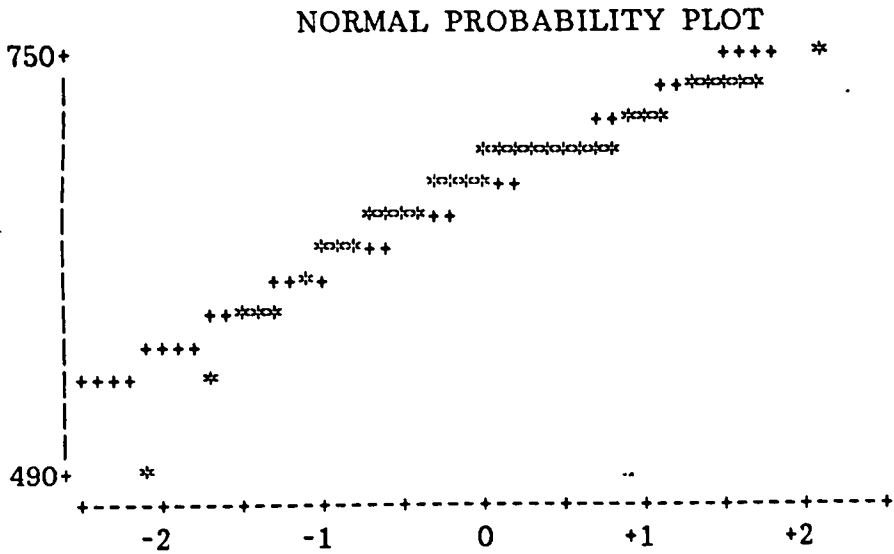
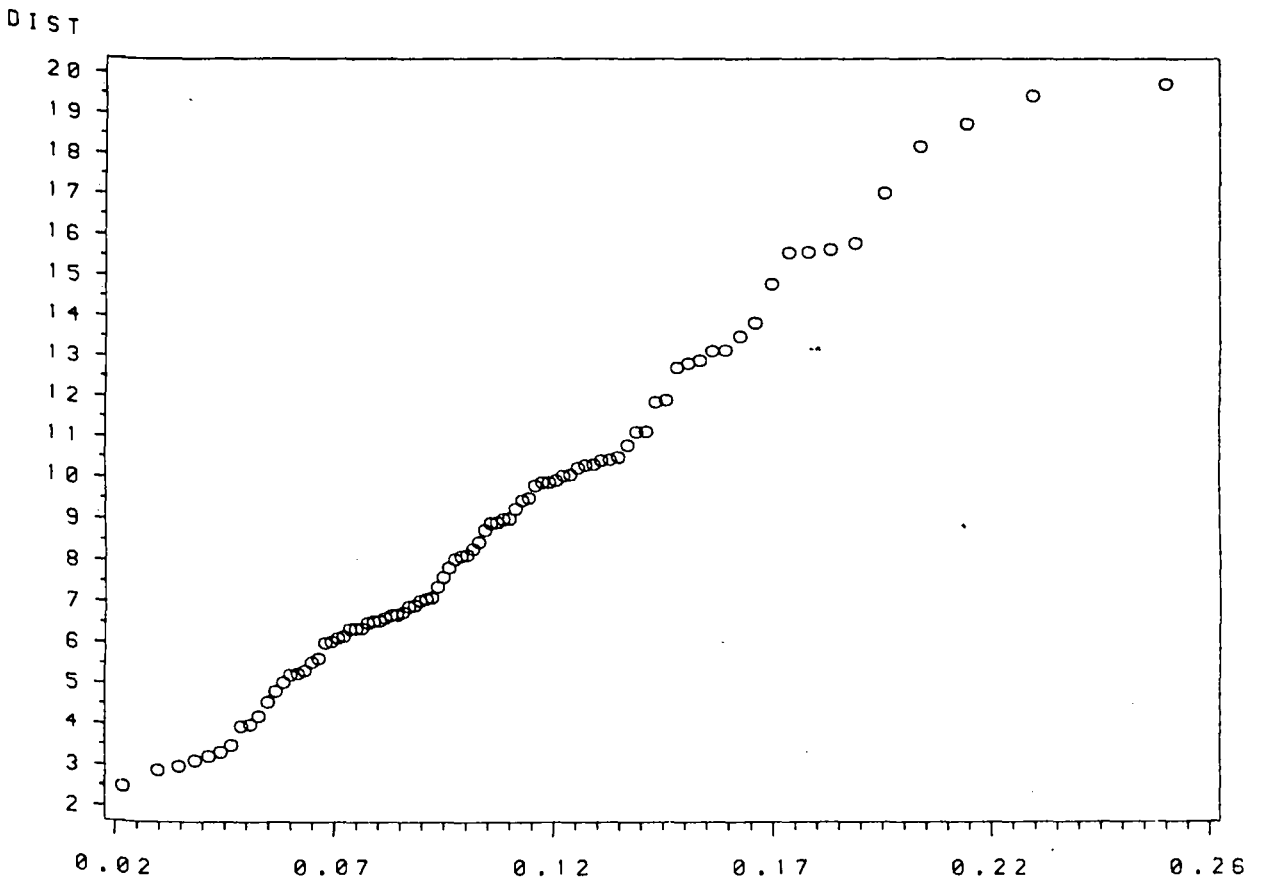
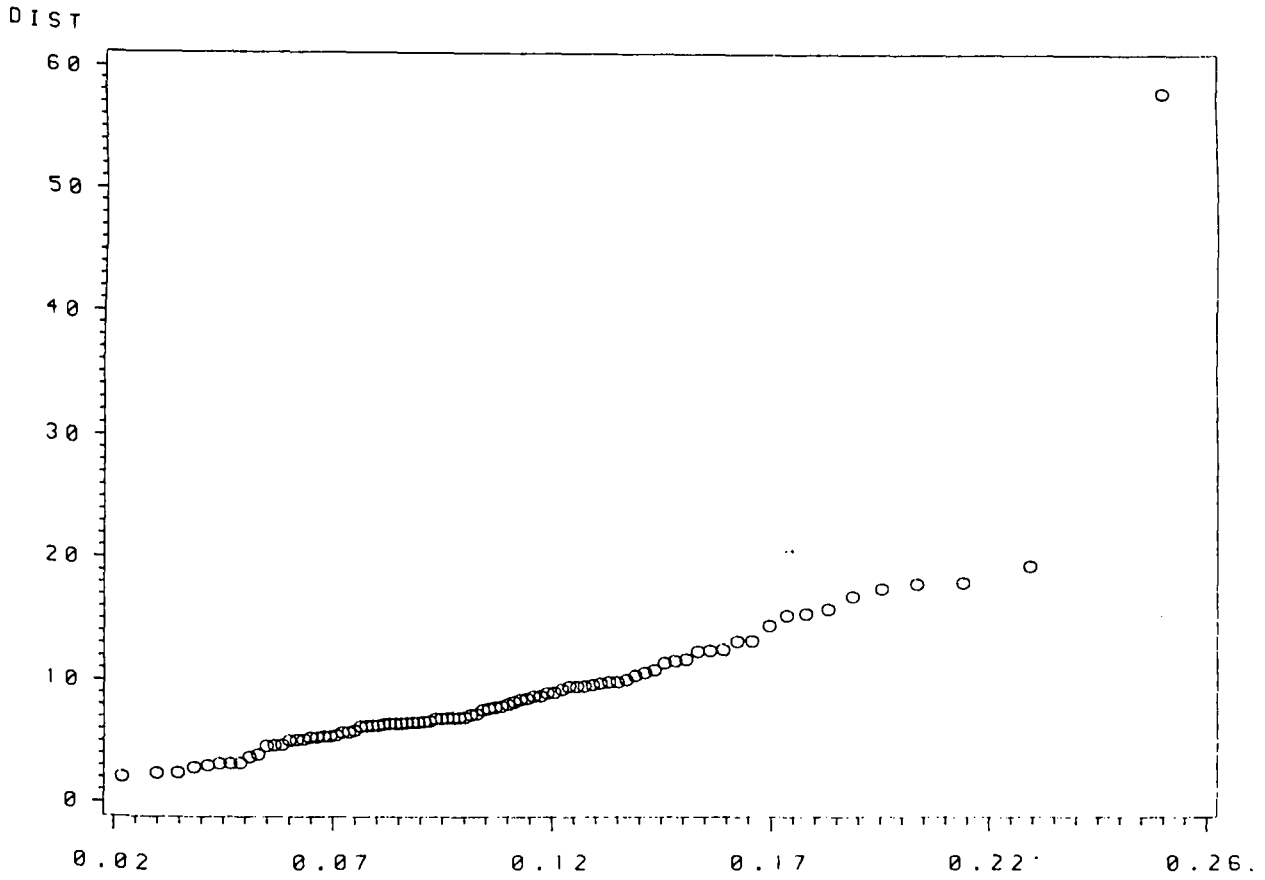


Figure 5.2

Q-Q plots of Mahalanobis' squared distances against quantiles of Beta distribution, before and after identification of outlier



CHAPTER SIX: WITHIN SPECIES VARIATION

6.1 INTRODUCTION

The analysis of variation within species using multivariate morphometric techniques is taxonomically important but inferentially very complex (Thorpe 1976, Blackith and Reyment 1971). Its importance lies in the deeper understanding of the biology of a species, which may have implications for its vectorial importance, but the complexity arises from difficulties in identifying the source and nature of the variation which has been uncovered.

The observed variation within a taxonomic unit is likely to be a combination of environmentally mediated variation and variation with a genetic basis, and can be expressed either as size differences, shape differences, differences in the covariation between characters, or any combination. This expressed variation could take the form of geographic and/or temporal variation between samples, and the patterns of variation could conform to any of a number of theoretical possibilities (Endler 1977).

In order to investigate fully intraspecific morphometric variation careful experimental planning and sampling regimes need to be adhered to. However, in the absence of carefully controlled experiment, valuable information can still be obtained from data acquired in an *ad hoc* manner. Thus the purpose of this chapter is to analyse, for the first time, multivariate morphometric variation within the *S. damnosum* complex, using data obtained primarily for purposes of discrimination between these species (Chapters seven and eight), to

describe this variation and to contrast it with known chromosomal or other variation, where appropriate.

6.2 MATERIALS AND METHODS

6.2.1 MATERIALS

The samples used in these analyses are listed as Appendix one, and were previously screened for outliers and/or contaminants (Chapter five). Of the 28 characters originally measured or scored, the number of spines on the basitarsus of the second leg was not analysed because it showed little interspecific variation and had many missing values (Chapter four). Two others, wing tuft colouration and abdominal setal colouration were analysed separately rather than jointly with the other characters because they both showed consistent departure from univariate normality (Chapter four).

6.2.2 DIMENSION REDUCTION

Even though three characters were not included in the analyses, 25 characters were measured or scored for each fly. This is clearly very large relative to most of the sample sizes obtained (Appendix one). Therefore, for the results not to be dominated by sampling error the number of characters jointly analysed needed to be reduced (Van Ness and Simpson 1976).

In the context of discriminant analysis, this is usually achieved by stepwise discriminant analysis (McKay and Campbell 1982a,b, Seber 1984). The pitfalls associated with this method are discussed in Chapter seven, where the inadequacy of the method in deriving a subset of characters with a low error rate is discussed. The analyses in the present chapter are not involved with the derivation of statistics for the future allocation of unknown flies in the field, but with describing intraspecific morphometric variation parsimoniously. In this context (i.e. descriptive discrimination rather than allocatory discrimination, Geisser 1977), stepwise discriminant analysis is more

appropriate for character subset generation, as the method optimises a measure of overall separation (Wilks' lambda) at each stage of the analysis (Seber 1984). The program used was SAS PROC STEPDISC, the algorithm that this program uses is described in section 7.2.1.

The initial subset of characters derived using stepwise discriminant analysis was examined for departures from univariate and multivariate normality, using SAS PROC UNIVARIATE for the marginal distributions and Mardia's multivariate skewness and kurtosis statistics for the joint distributions (see Chapter five). The null hypothesis of equal dispersion was also tested at this stage using the likelihood ratio test (see Chapter seven).

Characters which contributed to the departure from the null assumptions of equal dispersion and normality were removed interactively, until a character set was formed which was acceptable.

6.2.3 MULTIVARIATE ANALYSES

The Mahalanobis' squared distance between the sample mean vectors was calculated as part of the procedure SAS PROC DISCRIM and tested for significance. A canonical variates analysis was performed, involving extracting the eigenvalues and eigenvectors from the matrix $W^{-1}B$, where W is the pooled within-samples SSQPR matrix and B is the between-samples SSQPR matrix. The standardised eigenvectors (Canonical variates) show the association between each character and each discriminant vector, while the proportion of variation accounted for by each canonical (discriminant) vector is given by the associated eigenvalue (canonical root). The number of canonical variates extracted is $\min(g-1, p)$, where g is the number of groups and p is the number of characters. Thus for two groups there is one canonical

variate, which is directly related to the linear discriminant function.

The first canonical variate is that linear combination of the original characters along which the ratio of between groups to within groups variation is a maximum (Campbell and Atchley 1981), with subsequent vectors accounting for progressively smaller proportions of variance. Thus, the plane defined by the first two canonical variates is the best two dimensional representation of discriminant space amongst all such planes (Gnanadesikan 1977), and was used in these analyses as a parsimonious representation of this space. The procedure used was SAS PROC CANDISC.

A multivariate analysis of variance (MANOVA) was performed (using SAS PROC GLM) to assess the degree of overall separation *via* Wilks' lambda, which is the ratio of the determinant of the between-samples SSQPR to the determinant of the total SSQPR (Seber 1984).

The influence of size variation on between samples variation was assessed by calculating the eigenvectors of the pooled within-samples correlation matrix. The first eigenvector (principal component) is usually a size vector in morphometric studies (Rao 1964), and the scores along this vector were introduced as a covariable in a multivariate analysis of covariance (MANCOVA). The effect on the individual canonical roots was used as an informal measure of the influence of size on discrimination. The pooled within-samples correlation matrix was calculated by a routine written in SAS PROC MATRIX, while the MANCOVA was calculated using SAS PROC GLM.

6.3 RESULTS AND DISCUSSION

6.3.1 SIMULIUM DAMNOSUM S.S

Three samples of this species were available for analysis (Appendix one), one from Sierra Leone (V1=10, N=28), one from Benin (V1=28, N=16) and one from Togo (V1=28, N=27).

A stepwise discriminant analysis on the 25 character set resulted in an initial subset of eleven characters :

[V5,V9,V12,V15,V16,V17,V21,V22,V23,V25,V27]

i.e one head, three antennal, three wing and four leg characters. Mardia's multivariate skewness and kurtosis statistics and examination of the marginal distributions revealed some departure from the assumption of multivariate normality. The characters responsible for this distortion were removed interactively until the five character subset:

[V9,V16,V22,V25,V27]

resulted which conformed to the assumption of multivariate normality. In addition, the likelihood ratio test for equality of dispersion was not rejected at $p < 0.0001$, so two of the basic assumptions of the multivariate linear model were met.

Table 6.1 shows the overall mean, pooled coefficient of variation, and proportion of variance among localities for the five character subset. The CVs are all very similar, but the proportion of variance among samples is lower for antennal length 1 than for the other characters.

Using the five character subset, the matrix of Mahalanobis' squared distances shown in Table 6.2 was obtained. All of these distances were significant at $p < 0.001$, but clearly the two eastern

samples (V1=10, V1=28) are closer to each other than either is to the western sample (V1=10).

Examination of the standardised canonical variates shown in Table 6.3 reveals that the first canonical variate, with a canonical root of only 0.9004 (accounting for 80% of total variance) loads strongest on two characters, wing length 1 and femur length 2. The group means along this vector were: [-1.15,0.83,0.7] for samples V1=10,11,28 respectively, so it discriminates principally the western sample from the eastern samples, as can also be seen from Figure 6.1. The second canonical variate, with a canonical root of only 0.224 is therefore probably not biologically significant, even though it is statistically significant (Campbell 1982). This vector discriminates between the two eastern samples (Figure 6.1).

The first principal component of the pooled within-groups correlation matrix was a size vector accounting for 74% of pooled within-groups variance, with coefficients,

$$[0.33,0.47,0.49,0.46,0.48]$$

When the scores along this vector were introduced into the model as a covariable the canonical roots fell from 0.9004,0.2243 to 0.7169,0.122. With size controlled in this way, there was still significant morphometric differentiation between the samples, but this was reduced.

To conclude, there is significant morphometric differentiation between the three samples of *S. damnosum* s.s., but this is not great when compared with the results from some other groups (e.g. see section 6.3.4), and includes a significant proportion of size variation.

One sample (V1=10) is separated from the other two by two years and about 1500 km, and this sample is morphometrically further from the other two samples. However, *S. damnosum s.s.* is known to migrate distances greater than 500 km, (Garms and Walsh 1987), so it is possible that the eastern and western samples are from the same gene pool. Therefore the source of the multivariate variation found among samples within this species cannot be attributed to any specific causal factor such as seasonal or geographic variation.

Table 6.1

Overall mean, coefficient of variation (CV), and proportion of among-samples variance ($s^2(A)$)

| Character | Mean | CV | $s^2(A)$ (%) |
|-------------------|----------------|------|--------------|
| Antennal Length 1 | 256.55 μ m | 4.71 | 6.01 |
| Wing Length 1 | 731.35 μ m | 4.51 | 27.61 |
| Femur Length 1 | 631.42 μ m | 4.47 | 14.66 |
| Tarsus Segment 2 | 165.93 μ m | 4.56 | 20.28 |
| Femur Length 2 | 663.23 μ m | 4.05 | 25.14 |

Table 6.2

Mahalanobis' squared distances between samples

| From sample | 10 | 11 | 28 |
|-------------|------|------|------|
| 10 | 0.0 | 4.44 | 3.68 |
| 11 | 4.44 | 0.0 | 1.53 |
| 28 | 3.68 | 1.53 | 0.0 |

all $p < 0.001$

Table 6.3

Standardised Canonical Variates for *S. damnosum* s.s.

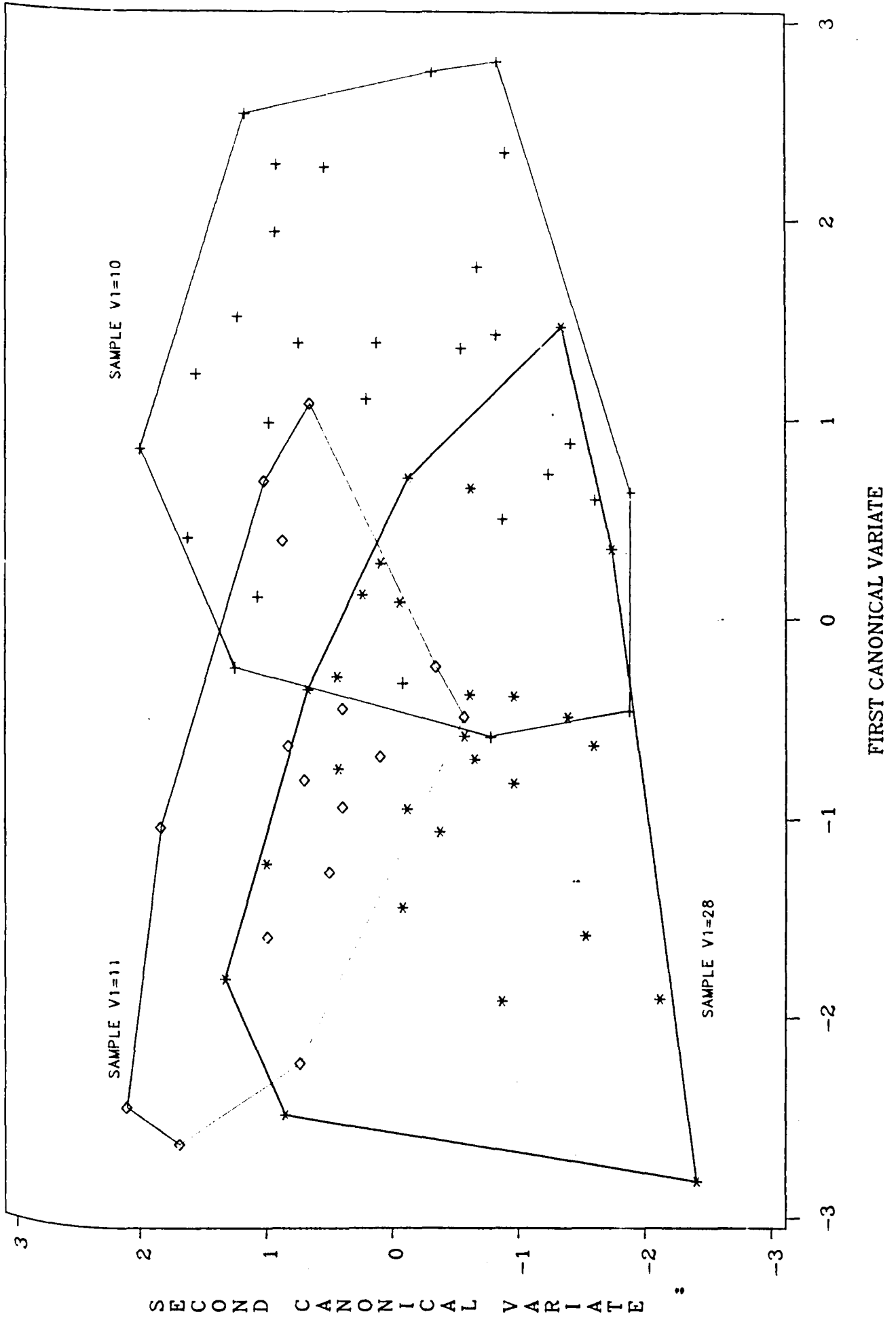
| Character | CV I | CV II |
|-------------------|---------------------|---------------------|
| Antennal Length 1 | -0.7088 | -0.3105 |
| Wing Length 1 | 1.2212 | -0.4399 |
| Femur Length 1 | -0.8888 | 0.5433 |
| Tarsus Segment 2 | -0.8196 | 1.4641 |
| Femur Length 2 | 1.4460 | -0.6032 |
| Canonical Root | 0.9004 ¹ | 0.2243 ² |

¹p<0.001

²0.001<p<0.01

Figure 6.1

Canonical Variates Plot for *S. damnosum* s.s.



6.3.2 SIMULIUM SIRBANUM

Five samples of *S. sirbanum* were available for analysis, four from Guinea (V1=21 N=44, V1=22 N=29, V1=23 N=43, V1=24 N=33), and one from Mali (V1=3 N=29) (Appendix one). A stepwise discriminant analysis on the 25 character set excluding wing tuft colouration, abdominal setal colouration, and basitarsal spine number resulted in an initial subset of 11 characters:

[V3, V4, V5, V6, V11, V17, V18, V20, V21, V23, V25]

i.e. two thorax, two head, one antennal, four wing and two leg characters. Examination of this character set using the methods described in Chapter five revealed some departure from multivariate normality. The characters having the greatest influence on non-normality were removed interactively, until a seven character subset which showed only minor departure from joint normality was obtained:

[V3, V4, V6, V17, V18, V20, V23]

i.e. two thorax, one head, three wing and one leg character. The likelihood ratio test for equality of dispersion was not rejected at $p < 0.0001$ so two of the assumptions of the multivariate linear model were not violated.

Table 6.4 gives the overall mean, coefficient of variation and proportion of variance among samples for the seven character subset. The proportion of variance among samples varies between characters, with wing width being considerably less variable among samples than the other characters. The two wing length characters show similar proportions of among samples variance as each other, implying that wing length varies among samples, while wing width varies less, i.e. wing shape differs among samples. When each sample's CVs were examined in greater detail using the variability profile method demon-

strated in Bird *et al.* (1981), it became clear that the characters differed in their variability within profiles among samples. Thus head width, wing width 1 and wing length 3 were significantly less variable than the other characters, and this relationship was consistent among samples (Friedman randomised block test, $p < 0.001$). Therefore wing shape is less variable than most other characters. A possible explanation for this might be that wing shape is under stronger natural selection for aerodynamic reasons.

Using this seven character set in a discriminant analysis, the matrix of Mahalanobis' squared distances given as Table 6.5 was obtained. Examination of this matrix shows that the principal morphometric differentiation within *S. sirbanum* involves the sample from Mali (V1=3) from the other four samples.

The standardised canonical variates are given as Table 6.6. The first two of the canonical variates are statistically significant, but the second canonical variate, with a canonical root of only 0.2579 is therefore unlikely to be of biological significance (Campbell 1982). Figure 6.2 shows the scatter of points in the first discriminant plane defined by the first two canonical variates. This figure clearly shows the differentiation of the Mali sample from the other four samples. The first canonical variate (Table 6.6) is principally a contrast between the wing characters wing width 1 and wing length 3, i.e. wing shape, with some influence from head width, confirming the result inferred from the univariate analyses.

The first principal component of the pooled within-species correlation matrix accounted for 85% of pooled within-species variation, and was a size vector with coefficients,

[.34, .39, .38, .37, .39, .38, .39].

When the scores along this vector were introduced into the model as a covariable the first canonical root fell from 1.1635 to 0.81101, showing that size has some importance in between-samples variation along this vector. The second canonical root fell from 0.2539 to 0.2432 showing that size is of no importance along this (biologically insignificant) vector.

To conclude, the major differentiation among the five samples of *S. sirbanum* involved one sample (V1=3) which is different from the other four in both shape and size. The main shape difference is wing shape, with the four Guinea samples having longer wings than the Mali sample, and all five having approximately the same wing width.

The Mali flies were sampled in November 1984, V1=21 and V1=22 in September 1986, V1=23 in August 1985 and V1=24 in December 1985, so that simple seasonal size variation can be discounted as the source of between samples variance. The Mali sample is geographically most isolated, but this species is known to migrate distances greater than 500 km (Garms and Walsh 1987), so all five samples are likely to be from the same gene pool. The only other difference between the Mali sample and the others is that it was reared from pupae rather than caught at human bait, so that the age distributions in the samples is likely to be different. Whether this is the source of variation, or some other unspecified source is responsible for the differentiation can not be decided on the available data.

Table 6.4

Overall mean, coefficient of variation (CV), and proportion of among-samples variance ($s^2(A)$)

| Character | Mean | CV | $s^2(A)$ (%) |
|----------------|-----------------|------|--------------|
| Thorax Length | 612.42 μ m | 6.8 | 19.4 |
| Thorax Width | 846.99 μ m | 6.42 | 19.17 |
| Head Width | 767.4 μ m | 5.37 | 28.43 |
| Wing Length 2 | 432.46 μ m | 6.35 | 25.95 |
| Wing Width 1 | 962.68 μ m | 5.6 | 4.19 |
| Wing Length 3 | 1394.33 μ m | 5.04 | 24.15 |
| Femur Length 2 | 657.29 μ m | 5.52 | 12.77 |

Table 6.5

Mahalanobis' squared distances between samples

| Sample | 3 | 21 | 22 | 23 | 24 |
|--------|-------------------|-------------------|-------------------|-------------------|------|
| 3 | 0.0 | 9.27 | 6.98 | 9.64 | 6.78 |
| 21 | 9.27 ¹ | 0.0 | 0.47 | 1.23 | 2.16 |
| 22 | 6.98 ¹ | 0.47 ³ | 0.0 | 1.41 | 1.25 |
| 23 | 9.64 ¹ | 1.23 ² | 1.41 ¹ | 0.0 | 0.99 |
| 24 | 6.78 ¹ | 2.16 ¹ | 1.25 ² | 0.99 ³ | 0.0 |

¹p<0.001

²0.01<p<0.05

³p>0.05

Table 6.6

Standardised Canonical Variates for *S. sirbanum*

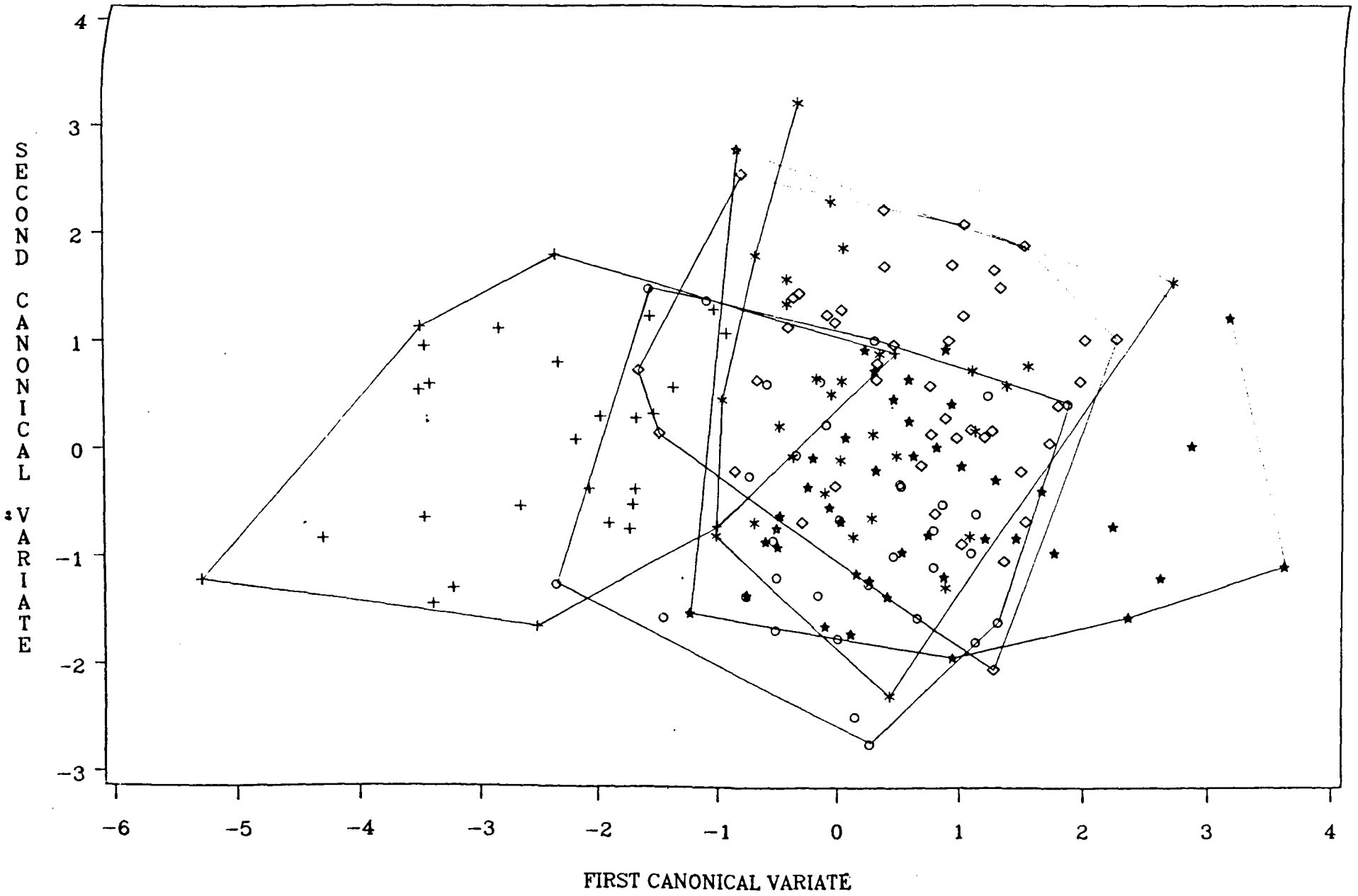
| Character | CVI | CVII | CVIII | CVIV |
|----------------|--------------------|--------------------|-------------------|--------------------|
| V3 | 0.37 | 0.31 | 0.99 | -0.35 |
| V4 | -0.29 | -2.45 | -0.27 | -0.79 |
| V6 | 1.09 | -0.5 | -0.76 | 1.8 |
| V17 | 0.89 | 0.10 | 0.32 | -0.59 |
| V18 | -1.85 | 0.05 | -0.41 | -0.37 |
| V20 | 1.63 | 1.4 | -1.2 | -1.24 |
| V23 | -0.91 | 0.87 | 1.8 | 1.42 |
| Canonical Root | 1.164 ¹ | 0.254 ¹ | 0.08 ² | 0.012 ² |

¹p<0.001

²p>0.05

Figure 6.2

Canonical Variates Plot for *S. sibiricum*



V1 + + + 3 ◇ ◇ ◇ 21 * * * 22 ★ ★ ★ 23 ○ ○ ○ 24

6.3.3 SIMULIUM SANCTIPAULI

Only two samples of this species were available for analysis (Appendix one), but these are chromosomally and geographically distinct. Sample V1=4 (N=35) is typical *S. sanctipauli* from Côte d'Ivoire, while sample V1=5 (N=26) is *S. sanctipauli* 'Djodji' form from Togo (Chapter two, Chapter three, Surtees *et al.* 1988).

A stepwise discriminant analysis between the two samples using the 25 character set resulted in an initial subset of eight characters:

[V5,V6,V9,V15,V16,V17,V19,V20]

however this character set showed departure from multivariate normality, using Mardia's skewness statistic and so was reduced to the two character subset:

[V9,V20]

i.e. antennal length and wing length. Table 6.7 gives the overall means, coefficient of variation and proportion of variance among samples for the two characters. The CVs are similar, but the proportion of variance among samples is strikingly different.

The Mahalanobis' squared distance between samples using the two characters was 3.33, which was significant at $p < 0.001$.

The standardised canonical variate, given as Table 6.8 was a shape vector, representing a contrast between antennal length 1 and wing length 3. Figure 3.3 shows the bivariate scatter of points using the two characters, and the differentiation between the two samples is clear.

The first principal component of the pooled within-samples correlation matrix was a size vector, with coefficients,

[0.71,0.71]

and accounted for 69.5% of pooled within-samples variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 0.8292 to 0.8279, indicating that the influence of size variation is negligible.

To conclude, there is significant morphometric differentiation between the two samples of *S. sanctipauli*. They were collected in the same month of the same year, so that simple temporality can be discounted as the source of the variation. The two samples do not differ along the size vector, so the observed morphometric variation is more likely to have a genetic basis. This variation is entirely shape variation involving the relative size of the antenna, a feature of *S. damnosum* s.l. morphology which has frequently been used to distinguish between species (e.g. Garms 1978), and which is presumably developmentally independent from the rest of the morphology. This species is known not to migrate large distances (Garms and Walsh 1987), allowing for the evolution of localised forms. Therefore, it appears that the chromosomal differentiation of *S. sanctipauli* 'Djodji' from typical *S. sanctipauli* has been paralleled by morphological evolution, although the extent of this will have to be established by more comprehensive sampling.

Table 6.7

Overall mean, coefficient of variation (CV), and proportion of among-samples variance ($s^2(A)$)

| Character | Mean | CV | $s^2(A)$ (%) |
|-------------------|-----------------|------|--------------|
| Antennal Length 1 | 328.21 μ m | 3.75 | 52.18 |
| Wing Length 3 | 1522.33 μ m | 4.13 | 2.73 |

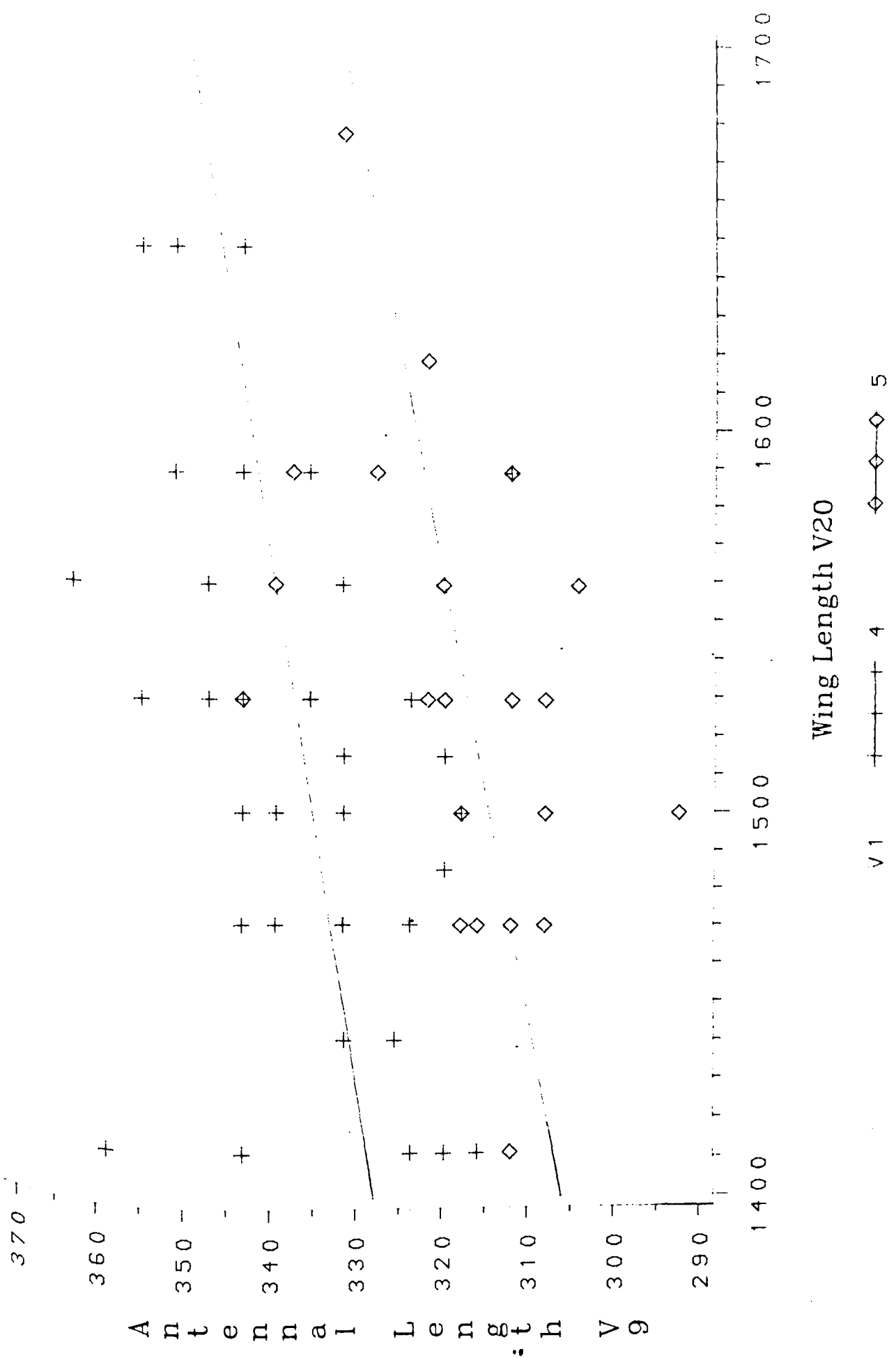
Table 6.8

Standardised Canonical Variate for *S. sanctipauli*

| Character | CVI |
|-----------|-------|
| V9 | 1.32 |
| V20 | -0.61 |

Figure 6.3

Scatter Plot of Wing Length against Antennal Length for *S. sanctipauli*



6.3.4 SIMULIUM SOUBRENSE AND S. SOUBRENSE 'B'

Eight samples were available within *S. soubrense* s.l., representing a considerable degree of chromosomal variation. One of these samples was *S. soubrense* 'B' (V1=6, N=28) which has been described as a separate species (Post 1986). However, it is chromosomally closest to *S. soubrense* (Chapter three), so to make the present analysis as general as possible, this sample was included in the analysis. In addition to this sample, there were two samples of *S. soubrense* 'Beffa' (V1=14 N=25, V1=29 N=11), one of *S. soubrense* 'Chutes Milo' (V1=16 N=30), three of *S. soubrense* 'Menankaya/Konkoure' (V1=17 N=32, V1=19 N=31, V1=25 N=27) of which V1=19 is chromosomally closest to *S. soubrense* 'Konkoure' *sensu* Quilléveré *et al.* 1982, V1=17 and V1=25 are closer to Menankaya form, see Chapter three). Finally one sample was of a form of unknown chromosomal affinities designated by Dr. R. Baker (personal communication) as *S. soubrense* 'Forest' form (V1=18 N=38).

Wing tuft colour was very variable between the samples of *S. soubrense*, with the null hypothesis of equal wing tuft colouration being rejected at $p < 0.001$ using a Kruskal-Wallis non-parametric one-way analysis of variance. Figure 6.4 shows the histograms for this character for the five subgroups within *S. soubrense* s.l.. The darkest flies belong to *S. soubrense* 'Forest' form, which confirms the original criterion for describing this morphological form. *S. soubrense* 'Beffa', *S. soubrense* 'B' and *S. soubrense* 'Chutes Milo' have similar frequency distributions which are basically symmetrical. The lightest of all the flies was *S. soubrense* 'Menankaya/Konkoure'. This result suggests that the chromosomal evolution within these species and forms has been paralleled by the expression of wing tuft

colouration. It was unfortunate that the standard scheme for scoring this character was adopted (Kurtak *et al.* 1981), as this was insensitive to the considerable taxonomic information in this character.

A stepwise discriminant analysis on the 25 character set resulted in an initial subset of 21 characters, which was too large for a parsimonious description of between-samples variation. Mardia's multivariate skewness and kurtosis statistics, together with examination of the marginal distributions and probability plotting of the joint distributions (Chapter five) resulted in a nine character subset:

[V4,V10,V17,V18,V20,V24,V27,V28,V29]

i.e. one thorax, one antennal, three wing and four leg characters. This character subset showed only moderate departures from multivariate normality, but the likelihood-ratio test for equality of dispersion was rejected at $p < 0.001$. Despite this, the pooled dispersion matrix was still used, for relative computational simplicity. The rejection of the likelihood ratio test for equal dispersion could be because of the relatively large number of characters, and the slight departure from normality, to which this test is not robust (Seber 1984).

Table 6.9 gives the overall means, pooled coefficient of variation, and proportion of among-samples variation for each character. The CVs suggest that antennal length and the two larger wing measurements might be less variable than the other characters. The proportion of among-samples variation is high for all characters, implying a considerable degree of morphological heterogeneity.

The nine character subset resulted in the matrix of Mahalanobis' squared distances shown as Table 6.10. All these distances were

significant at $p < 0.001$. As it stands this matrix is difficult to interpret, so a hierarchical cluster analysis method was used to simplify the matrix. The method used was the group average method (UPGMA, see Chapter three). The cophenetic correlation coefficient was only 0.58 so that the results of the cluster analysis should be treated cautiously.

The dendrogram resulting from the cluster method is shown as Figure 6.5. The tightest cluster on this dendrogram includes the *S. soubrense* 'B' sample and *S. soubrense* 'Chutes Milo' form, with a Mahalanobis' squared distance of 3.27 (significant at $p < 0.001$). The next tightest cluster contains the two *S. soubrense* 'Menankaya' samples, followed by a cluster consisting of the two *S. soubrense* 'Beffa' samples. Beyond these clusters, the inadequacy of the hierarchical representation allows only a limited interpretation, however, the two-cluster solution consists of one cluster containing all the western samples and the other cluster consisting of *S. soubrense* 'Beffa'. *S. soubrense* 'Forest' joins the western cluster last, and may therefore represent a morphologically distinct taxon relative to other *S. soubrense*.

The standardised canonical variates (Table 6.11) show the character combinations of importance in between-samples variation. The scatter of points in the first discriminant plane are shown as Figure 6.6. The first canonical variate has a canonical root of 2.1693, and is a contrast between two positively loading characters, wing width 1 and wing length 3, and two negatively loading characters, basitarsus length 1 and femur length 2, and therefore expresses a relationship between wing size and leg size (i.e. shape). This vector discriminates, at the positive end, the *S. soubrense* 'Menankaya/Konkoure'

samples, while at the negative end there is a sample of *S. soubrense* 'Beffa'. The other samples lie intermediate between these samples. The second canonical variate has a canonical root of 1.5858, and contrasts two positively loading characters: antennal length 2 and basitarsus length 2 against two negatively loading characters: wing width 1 and tibia length 2. This vector discriminates principally between *S. soubrense* 'Forest' and *S. soubrense* 'Konkoure'.

The first principal component of the pooled within-samples correlation matrix was a size vector, with coefficients,

[0.33,0.28,0.33,0.33,0.34,0.34,0.35,0.35,0.35]

and accounted for 83.7% of pooled within-samples variation. When the scores along this vector were introduced into the model as a covariable, the canonical roots fell from,

[2.1693,1.5858,0.5681,0.3244,0.1095,0.0964,0.023] to,

[1.829,1.2055,0.3509,0.1626,0.1051,0.041,0.0229].

Therefore, the first two canonical variates are both influenced a little by size variation, but discrimination between samples is still effective when size is controlled.

To conclude, there is extensive morphometric variation within *S. soubrense*/*S. soubrense* 'B'. This differentiation reflects the chromosomal heterogeneity of this group (Post 1986, Chapter three) but does not exactly parallel it. Thus, *S. soubrense* 'B' is chromosomally very distinct but morphologically it is very similar to other *S. soubrense*. The migratory ability of *S. soubrense* is not well known (Garms and Walsh 1987), but the considerable chromosomal and morphological heterogeneity within *S. soubrense* *s.l.* supports the hypothesis that it does not migrate far, allowing for localised differentiation into new chromosomal and morphological forms.

Table 6.9

Overall mean, coefficient of variation (CV), and proportion of among-samples variance ($s^2(A)$)

| Character | Mean | CV | $s^2(A)$ (%) |
|-------------------|-----------------|------|--------------|
| Thorax Width | 872.95 μ m | 5.48 | 35.96 |
| Antennal Length 2 | 452.31 μ m | 4.66 | 49.84 |
| Wing Length 2 | 443.02 μ m | 5.77 | 46.7 |
| Wing Width 1 | 997.16 μ m | 4.92 | 45.82 |
| Wing Length 3 | 1478.65 μ m | 4.79 | 55.23 |
| Basitarsus Length | 432.3 μ m | 5.31 | 43.75 |
| Femur Length 2 | 659.01 μ m | 5.01 | 38.48 |
| Tibia Length 2 | 617.97 μ m | 5.11 | 37.49 |
| Basitarsus Length | 325.77 μ m | 5.30 | 46.15 |

Table 6.10

Mahalanobis' squared distances between samples

| Sample | 6 | 14 | 16 | 17 | 18 | 19 | 25 | 29 |
|--------|--------------------|--------------------|-------------------|--------------------|--------------------|--------------------|-------|-----|
| 6 | 0.0 | | | | | | | |
| 14 | 10.04 ¹ | 0.0 | | | | | | |
| 16 | 3.27 ¹ | 11.16 ¹ | 0.0 | | | | | |
| 17 | 9.85 ¹ | 25.95 ¹ | 9.07 ¹ | 0.0 | | | | |
| 18 | 6.82 ¹ | 14.86 ¹ | 7.53 ¹ | 12.49 ¹ | 0.0 | | | |
| 19 | 6.87 ¹ | 17.6 ¹ | 6.51 ¹ | 6.98 ¹ | 18.46 ¹ | 0.0 | | |
| 25 | 10.72 ¹ | 24.21 ¹ | 8.0 ¹ | 4.66 ¹ | 7.26 ¹ | 12.7 ¹ | 0.0 | |
| 29 | 8.52 ² | 5.81 ³ | 8.23 ² | 10.71 ² | 10.0 ² | 10.68 ¹ | 11.22 | 0.0 |

¹p<0.001

²0.001<p<0.01

³p>0.05

Table 6.11

Standardised Canonical Variates for *S. soubrense* s.l.

| Character | CV I | CV II | CV III | CV IV | CV V | CV VI | CV VII |
|----------------|--------------------|--------------------|--------------------|--------------------|-------------------|--------------------|--------------------|
| V4 | -0.43 | 0.16 | 0.17 | 0.75 | 1.66 | -1.23 | -0.75 |
| V10 | 0.01 | 1.51 | -0.59 | -0.88 | -0.32 | 0.23 | 0.47 |
| V17 | 0.95 | -0.59 | 1.09 | -0.99 | -0.05 | -0.63 | 1.37 |
| V18 | 1.17 | -1.26 | -0.61 | -0.01 | -1.1 | 2.08 | -0.08 |
| V20 | 2.55 | 0.85 | -0.74 | -0.47 | 1.04 | -1.87 | -1.2 |
| V24 | -1.2 | 0.38 | 1.14 | -0.74 | -1.1 | 0.53 | -1.1 |
| V27 | -1.7 | -0.19 | 1.86 | -2.1 | 2.59 | 0.98 | -1.98 |
| V28 | -0.34 | -1.22 | -0.91 | 1.63 | -0.94 | -0.31 | 2.58 |
| V29 | 0.09 | 1.3 | -0.73 | 2.86 | 1.23 | 0.47 | 0.83 |
| Canonical Root | 2.169 ¹ | 1.586 ¹ | 0.568 ¹ | 0.324 ¹ | 0.11 ¹ | 0.096 ² | 0.023 ³ |

¹p<0.001²0.001<p<0.01³p>0.05

Figure 6.4

Histograms of Wing Tuft Colour for *S. soubrense* s.l.

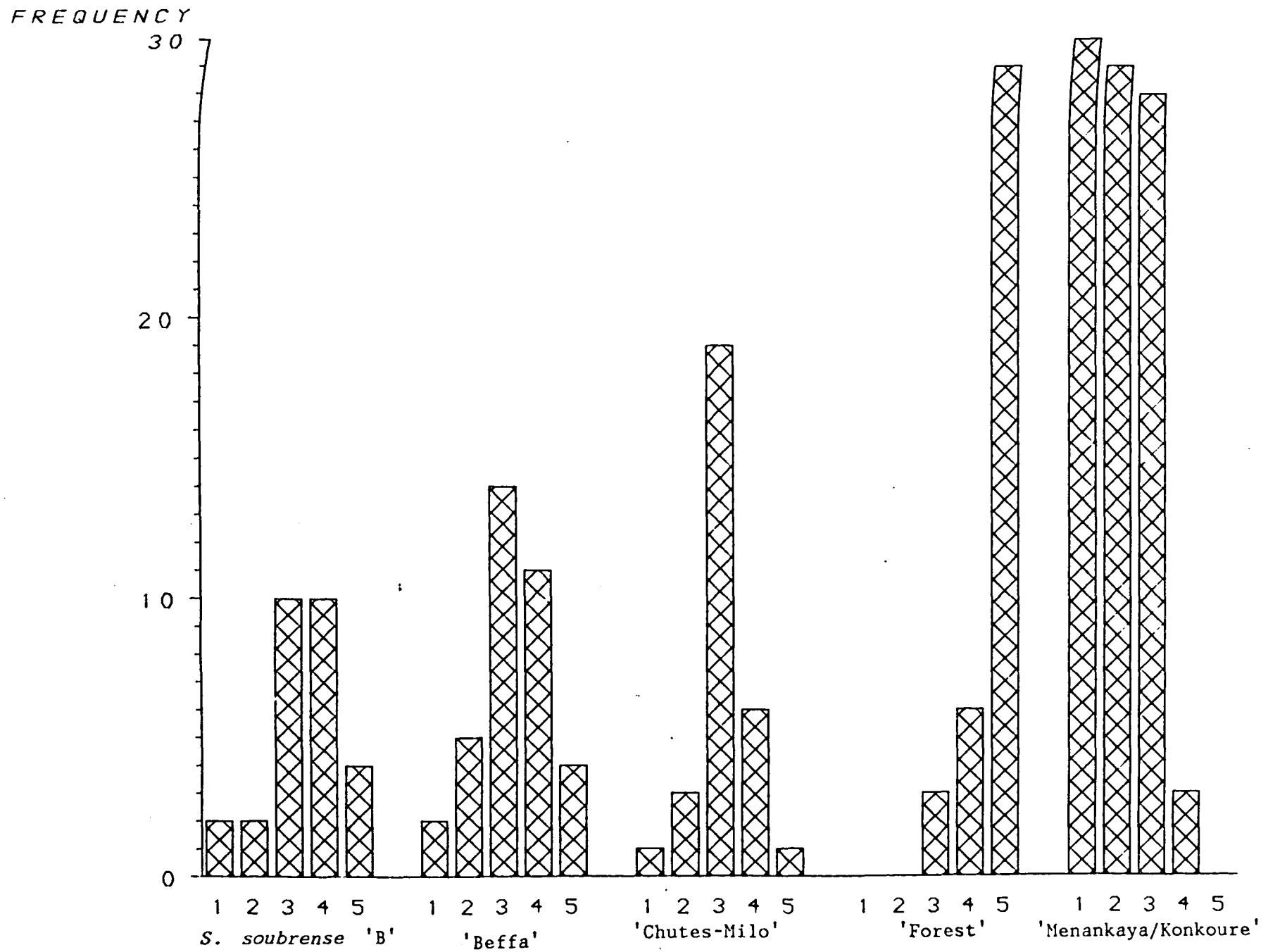


Figure 6.5

Dendrogram resulting from hierarchical cluster analysis of Mahalanobis' distance matrix between samples of *S. soubrense*

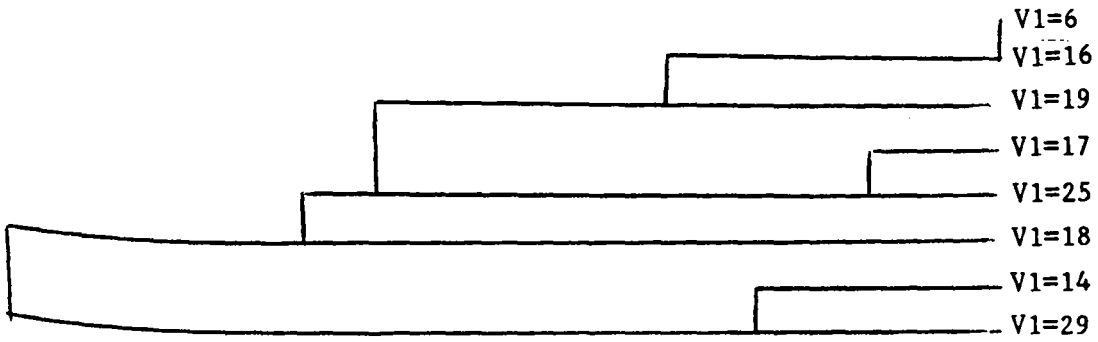
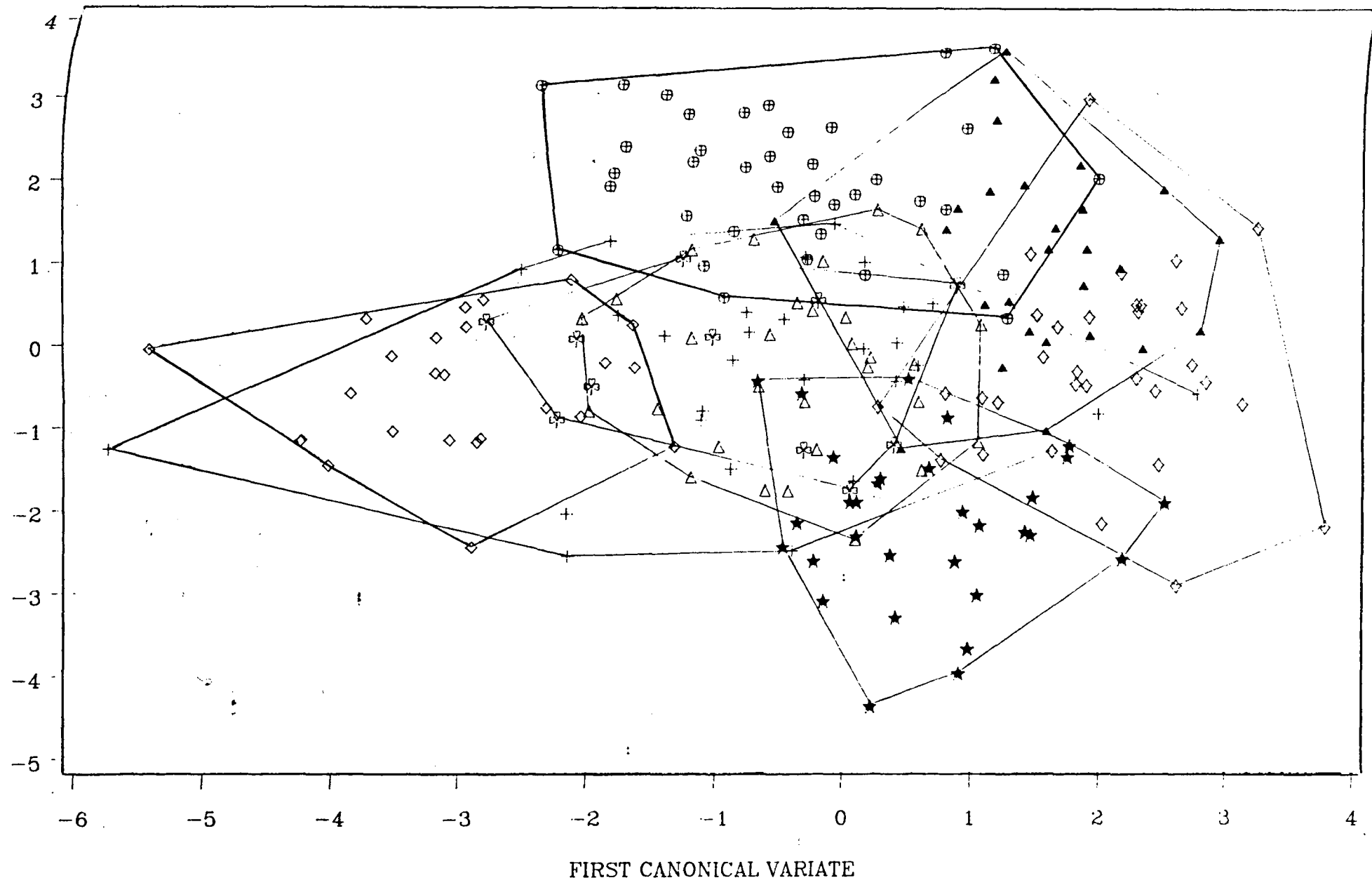


Figure 6.6

Canonical Variates Plot for *S. soubrense*

SECOND CANONICAL VARIATE



| | | | | | | | | | | | | | | | | |
|----|---|---|---|----|---|---|---|----|---|---|---|----|---|---|---|----|
| V1 | + | + | + | 6 | ◇ | ◇ | ◇ | 14 | △ | △ | △ | 16 | ◇ | ◇ | ◇ | 17 |
| | ⊕ | ⊕ | ⊕ | 18 | ★ | ★ | ★ | 19 | ▲ | ▲ | ▲ | 25 | ⊠ | ⊠ | ⊠ | 29 |

6.3.5 SIMULIUM SQUAMOSUM

Seven samples of *S. squamosum* were available (Appendix one), four from Togo (V1=1 N=34, V1=12 N=14, V1=13 N=20, V1=30 N=17) and three from Guinea (V1=15 N=40, V1=20 N=34, V1=26 N=13). Chromosomally the *S. squamosum* from the two countries differ, with the eastern *S. squamosum* showing less polymorphism than the western *S. squamosum* (Surtees and Post unpublished data), there is also evidence for allozymic differences between eastern and western *S. squamosum* (Townson *et al.* 1987) and DNA probes differences (Post personal communication).

A stepwise discriminant analysis on the 25-character set resulted in an initial subset of 13 characters:

[V4,V6,V9,V11,V12,V17,V19,V20,V21,V22,V26,V28,V29]

however, this subset showed some departure from multivariate normality and so was further reduced to the seven character subset:

[V9,V11,V19,V20,V22,V28,V29]

i.e. two antennal, two wing, and three leg characters.

The likelihood ratio test for equality of dispersion was not rejected at $p < 0.0001$.

Table 6.12 gives the overall means, pooled coefficients of variation and proportion of variance among samples. Some of the among-groups proportions of variance are very large implying considerable morphological differentiation. Wing length tends to be a less variable character, while antennal segment 4 is comparatively very variable, although this may be an artifact resulting from the very small size of this character (Lande 1977).

The matrix of Mahalanobis' squared distances between samples is given as Table 6.13. The most striking feature of this matrix is the

proximity of two of the Togolese samples ($V1=1$ and $V1=13$, $D^2=0.59$) but the distance of this pair from all other samples, including $V1=12$ which was sampled from the same site on other date as sample $V1=1$.

To examine this in further detail, the Togolese samples were analysed separately. Table 6.14 gives the matrix of Mahalanobis' squared distances between the Togolese samples using the seven character subset. Some of these distances are very large, and suggest considerable multivariate morphometric variation between samples of Togolese *S. squamosum*. However, when the scores along the first principal component of the pooled within-samples correlation matrix (which was a size vector accounting for 55.5% of pooled within-samples variation) were introduced into the model as a covariable, the first canonical root fell from 3.704 to 0.2193. Thus the observed variation within Togolese *S. squamosum* is size variation. This size variation is seasonal in origin, as the principal difference is between the pair: $V1=1$, $V1=13$ (sampled in October 1984 and October 1985 respectively) and $V1=12$ (sampled in March 1985). $V1=30$ was also sampled in October 1984 and is closest to the other October samples.

The western samples were also analysed separately for comparison, and the matrix of Mahalanobis' distances is shown as Table 6.15. The canonical roots changed from (0.935,0.272) to (0.695,0.24) when the scores along the first principal component of the pooled within-samples correlation matrix were introduced as a covariable, showing that size has some influence on variation between samples of Guinean *S. squamosum*.

The Mahalanobis' squared distance between the eastern *S. squamosum* and the western *S. squamosum* was 4.55 and the canonical root 1.1509. When size was controlled for by the introduction of the scores along

the first principal component of the pooled within-samples correlation matrix, the canonical root fell to 0.638, showing that there was little east/west differentiation beyond that already found within these areas.

Table 6.16 gives the standardised canonical variates for the eastern and western samples combined. The first canonical variate is a contrast between wing length 3 and tibia length 2/basitarsus length 2. The samples V1=1 and V1=13 are the main samples to be discriminated along this vector and when size is controlled for, the first canonical root fell from 2.35 to 0.927, indicating that the first canonical variate is dominated by the influence of size variation.

To conclude, there is a large amount of seasonal size variation within Togolese *S. squamosum*. Cheke and Harris (1980) recorded seasonal size variation in *S. damnosum s.l.* from Côte d'Ivoire and Cheke and Denke (1988) also noted seasonal size variation in *S. squamosum* from Togo, so the phenomenon is not unknown in *S. damnosum s.l.* Unfortunately all of the western samples were collected in the same month of 1986, so it is not possible to establish whether the extensive size variation is a species specific phenomenon, or whether it is a feature only of the eastern *S. squamosum*.

There was no evidence for morphometric differentiation between eastern and western *S. squamosum*, beyond that which was found within countries, once size was controlled for. Wing tuft colour showed some tendency to be differently distributed in the east than the west, with the former being lighter. Garms and Walsh (1987) suggest that *S. squamosum* in the West might be isolated from that in Togo, but morphology does not support the evidence for separate gene pools that

is implied by other sources of variation, including chromosomal, biochemical and ecological variation.

Table 6.12

Overall mean, coefficient of variation (CV), and proportion of among-samples variance ($s^2(A)$)

| Character | Mean | CV | $s^2(A)(\%)$ |
|--------------------|----------------------|------|--------------|
| Antennal Length 1 | 279.51 μm | 5.11 | 31.8 |
| Antennal Segment 4 | 37.7 μm | 6.83 | 44.05 |
| Wing Width 2 | 1466.4 μm | 4.73 | 56.44 |
| Wing Length 3 | 1540.1 μm | 4.28 | 56.21 |
| Femur Length 1 | 654.56 μm | 4.76 | 65.65 |
| Tibia Length 2 | 637.4 μm | 4.71 | 65.88 |
| Basitarsus Length | 343.4 μm | 4.77 | 66.2 |

Table 6.13

Mahalanobis' squared distances between samples

| Sample | 1 | 12 | 13 | 15 | 20 | 26 | 30 |
|--------|--------------------|--------------------|-------------------|-------------------|-------------------|-------------------|-----|
| 1 | 0.0 | | | | | | |
| 12 | 13.21 ¹ | 0.0 | | | | | |
| 13 | 0.59 ³ | 12.13 ¹ | 0.0 | | | | |
| 15 | 10.17 ¹ | 8.01 ¹ | 11.47 | 0.0 | | | |
| 20 | 14.9 ¹ | 2.84 ¹ | 15.33 | 4.76 ¹ | 0.0 | | |
| 26 | 8.2 ¹ | 3.64 ² | 9.48 ¹ | 4.14 ³ | 3.06 ³ | 0.0 | |
| 30 | 3.59 ² | 4.96 ¹ | 3.91 ¹ | 6.74 ¹ | 5.17 ¹ | 2.23 ² | 0.0 |

¹ $p < 0.001$

² $0.001 < p < 0.01$

³ $p > 0.05$

Table 6.14

Mahalanobis' squared distances between Togo samples

| Sample | 1 | 12 | 13 | 30 |
|--------|-------------------|-------|------|-----|
| 1 | 0.0 | | | |
| 12 | 26.93 | 0.0 | | |
| 13 | 0.62 ¹ | 23.01 | 0.0 | |
| 30 | 5.84 | 9.36 | 4.81 | 0.0 |

¹ $p > 0.05$

Table 6.15

Mahalanobis' squared distances between Guinea samples

| Sample | 15 | 20 | 26 |
|--------|------|------|-----|
| 15 | 0.0 | | |
| 20 | 4.12 | 0.0 | |
| 26 | 3.72 | 2.72 | 0.0 |

Table 6.16

First Three Standardised Canonical Variates for *S. squamosum*

| Character | CV I | CV II | CV III |
|----------------|-------------------|-------------------|-------------------|
| V9 | -0.37 | 0.09 | -0.93 |
| V11 | 0.45 | -0.29 | -0.12 |
| V19 | 0.43 | -1.67 | -0.49 |
| V20 | -1.16 | 1.93 | -1.3 |
| V22 | 0.39 | 2.13 | 2.64 |
| V28 | 0.91 | -1.32 | -0.83 |
| V29 | 1.09 | -0.63 | 0.27 |
| Canonical Root | 2.35 ¹ | 0.74 ¹ | 0.16 ¹ |

¹p<0.001

²0.001<p<0.01

6.3.6 SIMULIUM YAHENSE

Four samples of *S. yahense* were available for analysis (Appendix one), one from Côte d'Ivoire (V1=2 N=37), two from Sierra Leone (V1=7 N=14, V1=8 N=25) and one from Guinea (V1=27 N=20). Chromosomally these are differentiated in that the Côte d'Ivoire *S. yahense* is not sex-linked for the inversion 2L-18 (it is fixed), while in Sierra Leone the inversion is X-linked (Vajime and Dunbar 1975, Surtees and Post unpublished data). This is the only chromosomal heterogeneity recorded for *S. yahense*.

A stepwise discriminant analysis resulted in an initial subset of 11 characters:

[V5, V6, V12, V15, V17, V18, V19, V20, V22, V27, V28]

This character set showed considerable departure from multivariate normality, so the data set was reduced further to give the following seven character subset:

[V6, V17, V19, V20, V22, V27, V28]

i.e. one head, three wing and three leg characters. However, this subset also showed multivariate skewness, and the likelihood ratio test for equality of dispersion was rejected at $p < 0.001$. Thus two of the assumptions of the model were rejected, meaning that interpretation of the results obtained should be cautious.

Table 6.17 gives the overall mean, coefficients of variation and proportion of among-samples variance for the seven characters. The CVs are comparable although there is some indication that head width and two of the wing characters are less variable than the other characters. This was supported by Friedman's test for randomised blocks ($p < 0.01$) applied to each sample's variability profile, showing that the profiles are consistent among samples (Bird *et al.* 1981).

Table 6.18 gives the matrix of Mahalanobis' squared distances between samples using the seven character subset. It is apparent that one of the Sierra Leone samples, V1=7 is differentiated relative to the other samples.

The standardised canonical variates are shown in Table 6.19. The first canonical variate is mainly a contrast between femur length 2 and tibia length 2 with some influence from wing length. This vector discriminates sample V1=2 at its positive end from V1=7 at the negative end. The second canonical variate only has a canonical root of 0.447 and is therefore not biologically significant (Campbell 1982). It contrasts two positive characters wing width 2 and wing length 3 with two negative characters: femur length 1 and femur length 2. This vector discriminates the sample V1=8 from the other samples. Figure 6.7 shows the scatter of points in the first principal plane.

The first principal component of the pooled within-samples correlation matrix was a size vector with coefficients,

$$[0.36, 0.36, 0.37, 0.38, 0.39, 0.40, 0.39]$$

and accounted for 86% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable, the canonical roots changed from [1.223, 0.447, 0.359] to [0.453, 0.364, 0.245] indicating that the first canonical variate was substantially influenced by size variation.

To conclude, there is significant multivariate morphometric variation between the samples of *S. yahense*, but this variation is principally size variation. There is no indication whether this variation is geographic or temporal in origin, and the chromosomally distinct Côte d'Ivoire sample is not morphologically distinct. This species is known to migrate only very short distances (Garms and Walsh

1987), but despite this there is no evidence for localised morphological differentiation.

Table 6.17

Overall mean, coefficient of variation (CV), and proportion of among-samples variance ($s^2(A)$)

| Character | Mean | CV | $s^2(A)(\%)$ |
|----------------|-----------------|------|--------------|
| Head Width | 828.68 μ m | 4.53 | 38.81 |
| Wing Length 2 | 471.04 μ m | 5.95 | 40.4 |
| Wing Width 2 | 1495.93 μ m | 4.76 | 47.44 |
| Wing Length 3 | 1576.96 μ m | 4.57 | 50.47 |
| Femur Length 1 | 657.64 μ m | 5.61 | 48.29 |
| Femur Length 2 | 690.34 μ m | 5.22 | 51.26 |
| Tibia Length 2 | 648.42 μ m | 5.59 | 40.81 |

Table 6.18

Mahalanobis' squared distances between samples

| Sample | 2 | 7 | 8 | 27 |
|--------|-------|------|------|-----|
| 2 | 0.0 | | | |
| 7 | 10.36 | 0.0 | | |
| 8 | 4.11 | 5.96 | 0.0 | |
| 27 | 4.32 | 5.93 | 3.63 | 0.0 |

Table 6.19

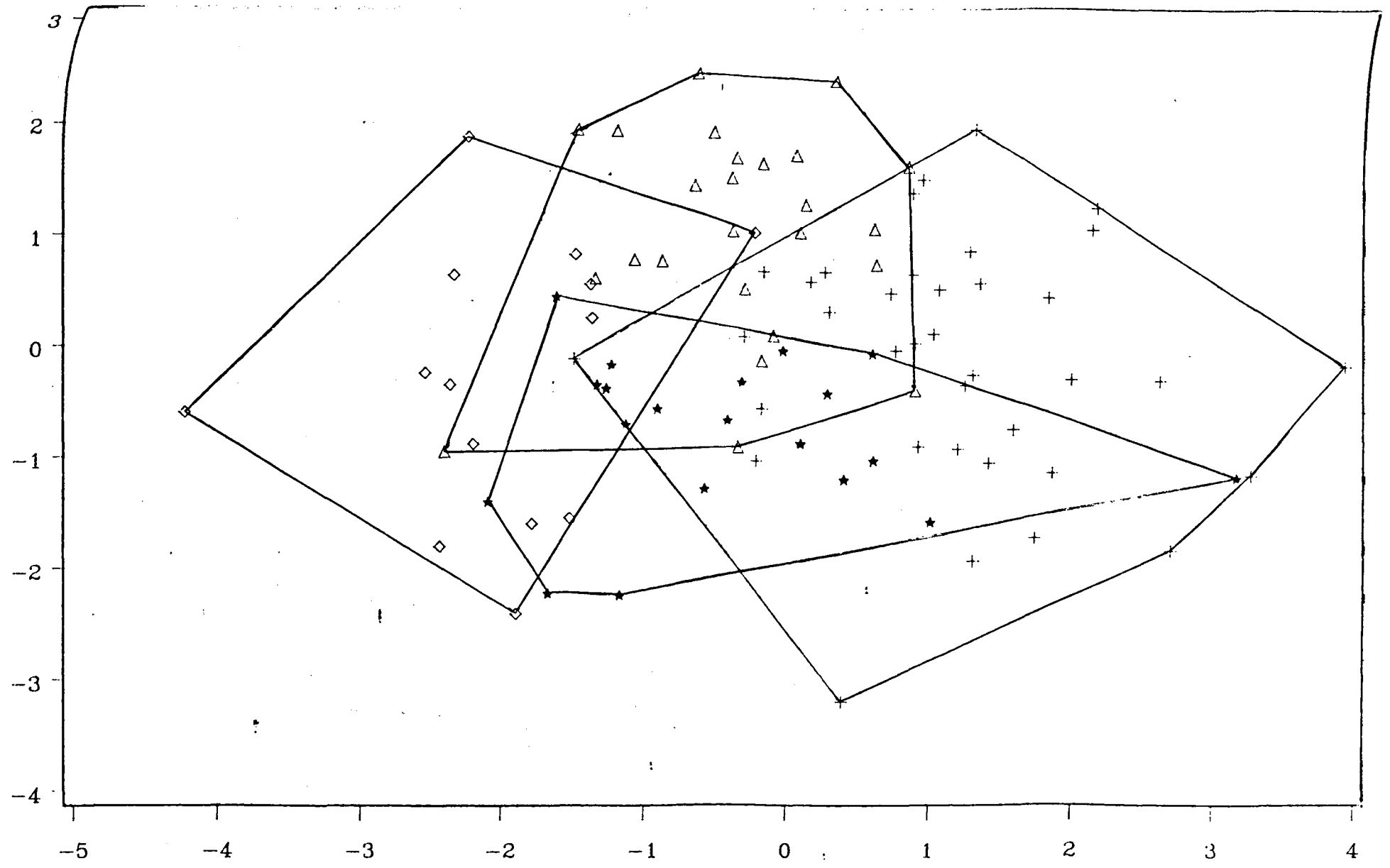
Standardised Canonical Variates for *S. yahense*

| Character | CV I | CV II | CV III |
|----------------|--------------------|--------------------|--------------------|
| V6 | -0.38 | -0.44 | -1.17 |
| V17 | 0.14 | 0.05 | -1.86 |
| V19 | 0.18 | 2.03 | 0.57 |
| V20 | 1.3 | 1.2 | -1.22 |
| V22 | -0.07 | -1.83 | 0.06 |
| V27 | 2.67 | -1.54 | 1.64 |
| V28 | -2.51 | 0.62 | 1.85 |
| Canonical Root | 1.223 ¹ | 0.447 ¹ | 0.359 ¹ |

¹p<0.001

Figure 6.7

Canonical Variates Plot for *S. yahense*



v1 + + + 2 ◇ ◇ ◇ 7 △ △ △ 8 ★ ★ ★ 27

6.4 DISCUSSION

Table 6.20 summaries the results presented in sections 6.3.1-6.3.6. The percentage of variation along the first principal component of the pooled within-samples correlation matrix (the size vector) differed between species, but was consistently high. This reflects the high intercorrelation between most of the characters used to describe variation within species.

The size-free canonical roots can be used as an informal indicator of the amount of shape differentiation within the six taxa. Clearly, *S. soubrense* shows the greatest morphological differentiation between samples, as even the second canonical root is larger than any other canonical root for the other species. The other species show similar degrees of morphological differentiation, which is generally not great, especially when compared with between-species analyses (see Chapters seven and eight).

The influence of size variation on among-samples variation differs greatly between species. *Simulium yahense* and *S. squamosum* are both very heavily influenced by size variation, as shown by the ratio of size-free canonical root to total canonical root, expressed as a percentage. For both of these species, among-samples size variation seems to be the predominant mode of variation. Whether this similarity between the two species is a reflection of their chromosomal relatedness can not be stated on the present data. *S. sirbanum*, *S. damnosum* s.s. and *S. soubrense* are decreasingly influenced by size variation among samples, while *S. sanctipauli* is uninfluenced by size variation. This result could be an artifact of the limited sampling of *S. sanctipauli*, and only additional data will clarify this.

Thus three sources of between-samples variation have been uncovered in the six main taxa of the *S. damnosum* complex. Temporal size variation is strongest in *S. squamosum*, and the extent of this variation is quite surprising. Geographic shape variation is most marked in *S. soubrense*, although this variation does not exactly parallel the considerable chromosomal variation within this taxon. Shape variation was less in *S. sanctipauli*, but may reflect the chromosomal differences between the two samples. The remaining species show varying degrees of size and shape variation, which cannot be attributed to any one specific cause.

To conclude, there is significant morphometric differentiation within the six main taxa of the *S. damnosum* complex. The extent of this varies between species, as does the influence of size variation.

Table 6.20

Summary of intraspecific morphometric analyses of the six main *S. damnosum* s.l. taxa

| Analysis ¹ | % Variation along first PC ² | Size-free canonical root | Size influence ³ |
|-----------------------|---|--------------------------|-----------------------------|
| 6.3.1 | 74% | 0.717, 0.122 | 79.6%, 53.4% |
| 6.3.2 | 85% | 0.811, 0.243 | 69.5%, 95.8% |
| 6.3.3 | 69.5% | 0.828 | 99.84% |
| 6.3.4 | 83.7% | 1.829, 1.206, 0.35 | 84.3%, 76%, 61.7% |
| 6.3.5 | 69.6% | 0.927, 0.348 | 39.5%, 46.87% |
| 6.3.6 | 86% | 0.453, 0.364 | 35.3%, 35.3% |

¹Numbers refer to analyses in section 6.3

²This is the proportion of variation along the first principal component of the pooled within-samples correlation matrix, expressed as a percentage

³This is the ratio of the size-free canonical root to the total canonical root expressed as a percentage. Only canonical roots considered of biological significance are included.

CHAPTER SEVEN: REGIONAL DISCRIMINATION OF FLIES

7.1 INTRODUCTION

The techniques of multivariate morphometric analysis can be used to study such areas as quantitative genetics, congruence of life stages in classification and geographic or other variation within species (Chapter six). However, the ultimate aim of the multivariate morphometric analysis of the *S. damnosum* complex is to produce a scheme for the allocation of unknown flies which is as accurate and as informative as possible.

The *S. damnosum* complex is very heterogeneous chromosomally (Vajime and Dunbar 1975, Post 1986, Chapter three) with particular variants only being found in certain restricted geographic areas. For example, in Togo there are the chromosomally distinct forms *S. soubrense* 'Beffa' (Meredith *et al.* 1983) and *S. sanctipauli* 'Djodji' (Surtees *et al.* 1988, Chapter two). *Simulium squamosum* in Togo is different chromosomally (Post and Surtees unpublished data), morphologically (Chapter six), and at the molecular level (Post personal communication), from *S. squamosum* further west in Sierra Leone. Within Guinea, there are the forms *S. soubrense* 'Menankaya/Konkoure', not found in Togo (Post personal communication), and from Sierra Leone and Guinea there is the species *S. soubrense* 'B' (Post and Crosskey 1985, Post 1986).

In view of this heterogeneity it was decided to derive the statistics for species identification from samples from Togo and Benin separately from the samples of species from Mali, Guinea, Sierra Leone

and Côte d'Ivoire. The purpose of this chapter is to describe the method used to derive the best character subsets from the full set of characters described in Chapter four, in the context of allocatory discriminant analysis and to present the results obtained using these character subsets, for the two geographic areas. In Chapter eight, a 'global' approach will be taken to the allocation of unknown flies which will not assume prior knowledge of the source of the fly beyond the fact that it is West African *S. damnosum s.l.*. All the methods described in this chapter were also used in Chapter eight.

In both chapters, statistics for pairwise discrimination of species are derived in addition to statistics for the allocation of an unknown fly to any of the reference species simultaneously. The pairwise statistics can be used preferentially in areas where it is known *a priori* that only two species are likely to occur (e.g. River Gban-Houa, Djodji, Togo; Garms and Cheke 1985, Surtees *et al.* 1988, Chapter two) or subsequent to the 'all-species' allocation using typicality probabilities, to improve allocation rates.

7.2 MATERIALS AND METHODS

7.2.1 MATERIALS

The full list of samples is given in Appendix one. These were all screened for outliers or split into their component species using the methods described in Chapter five. Of the 28 characters described in Chapter four, basitarsal spine number was not included in any of the analyses to follow for the reasons given in Chapter five.

7.2.2 SELECTION OF CHARACTER SUBSETS FOR DISCRIMINATION

A major factor which determines the usefulness of an allocation scheme derived using multivariate morphometric analysis is the number

of characters which need to be measured or scored for each fly. This analysis began with 28 characters (Chapter four). Such a large character set reduces the usefulness of the method for the following reasons:

1. It takes longer to identify a fly if all characters need to be measured or scored.
2. The probability that a fly will be complete for all the characters is smaller the larger the number of characters.
3. The estimates of the joint (multivariate) distribution are less accurate for a fixed sample size if the dimensionality (number of characters) is increased (Van Ness and Simpson 1976), which in practise has the effect of increasing the observed error rate (Van Vark 1984).
4. The assumptions of multivariate normality and equality of dispersion are less likely to be met for a larger character set than for a smaller one (Seber 1984).
5. More data must be stored in computer memory and more computation is involved in the field identification of a fly.

For these reasons it was necessary to find a method for finding a subset of characters which performed as well as the original full character set, in the context of allocatory discriminant analysis.

There are a number of methods available for character subset generation (Seber 1984), either within widely available computer packages, or as programs not available to this analysis (e.g. ALLOC80, Hermans *et al.* 1982), or as published algorithms (e.g. McKay and Campbell 1982a). For this analysis the only methods available were those within the statistical packages SAS and SPSSX.

The commonly available methods for deriving character subsets in the context of multiple group allocatory discriminant analysis are known to be inadequate (Seber 1984, Habbema and Hermans 1977, McKay and Campbell 1982b). The best of the widely available methods (stepwise discriminant analysis) is known to separate groups which are already well separated at the expense of groups which are close together, and may result in a character subset which performs poorly in an allocation scheme (Habbema and Hermans 1977). This is because the method optimises Wilks' lambda (a statistic describing overall discrimination) at each step in the analysis rather than optimising some statistic relevant to future allocation, such as error rate.

Because programs using the preferred methods (concentrating on error rate), were not available in this analysis, the following method was used for the generation of character subsets:

1. Stepwise discriminant analysis was used to generate an initial subset of characters (excluding wing tuft colouration and abdominal setal colouration because of their consistent non-normality, Chapter four), using the SAS PROC STEPDISC program. This method starts with a model containing no characters. At each step, characters in the model are tested, using Wilks' ratio (see Seber 1984), and any which fall below a preset constant (F-to-leave, set to 0.15, Costanza and Afifi 1979) are removed. Characters not in the model are then tested, and the one which best exceeds another preset constant (F-to-enter, also set to 0.15) is entered. The method stops when none of the characters outside of the model exceed F-to-enter and none of the characters in the model fall below F-to-leave.

The error rate resulting from this method was compared with that using the full 25 character set using the 'leave-one-out' method of

error rate estimation (Lachenbruch and Mickey 1968, also known as cross-validation and jackknifing, Seber 1984) for small to medium sized data sets (this error rate was calculated using an inefficient GENSTAT macro which could work only on moderately sized data sets), or, more usually, using the resubstitution method of error rate estimation (using SAS PROC DISCRIM). The 'leave-one-out' method works by removing a fly from the data set, calculating the discriminant functions in its absence and classifying it using the resultant statistics. Resubstitution works by classifying each fly in the data set using discriminant functions derived from the whole data set, including the fly to be classified. This method is more prone to bias than the 'leave-one-out' method, especially for small data sets and large character sets (Seber 1984). If the number of characters was sufficiently low and the error rate acceptable, then this initial subset was accepted. Otherwise:

2. Characters were removed one-by-one from the model, starting with those which had entered it last in the stepwise discriminant analysis. If the removal of the character had a detrimental effect on error rate, then it was returned into the model, however, if its removal had only a slight effect on error rate, then it was rejected. The process was stopped when all of the characters remaining in the model had a detrimental effect on error rate when removed.

As it stands, this method is open to some subjectivity in the choice of a particular subset of characters, because error rate was not the only parameter influencing character choice. For example, if a character increased error rate when it was removed, but it was known to depart from normality, or it was a relatively difficult

character to measure (Chapter four), then it was still rejected in spite of the detrimental effect on error rate.

Whilst it is difficult to demonstrate that the subsets of characters obtained by this method are optimal, the results which follow show that they are at least adequate.

7.2.3 PRIOR PROBABILITIES AND NON-NORMAL CHARACTERS

The previous section described the method used for the selection of subsets of characters. This method excluded two characters, wing tuft colouration and abdominal setal colouration because both have previously been shown to be non-normally distributed (Chapter four). However, these characters were included in the analysis because both have been shown to be of considerable taxonomic importance (Dang and Peterson 1980, Garms and Zillman 1984). The purpose of this section is to describe a method for including the taxonomic information these characters contain in the analyses to follow.

The simplest approach to including these characters in the analysis would be to derive the character subset of approximately normally distributed characters using the method described in the previous section and then to introduce the two colour characters into the linear discriminant functions (LDFs), keeping them in the analysis if error rate is improved. However there are two objections to this approach, one of which will be explained in this section, the other in the next section.

The first objection concerns the robustness of the LDF to departures from the assumptions of the model (the most important of which are multivariate normality and equality of dispersion). The LDF is generally robust to departures from normality (Lachenbruch 1975), however it is particularly sensitive to skewed continuous distrib-

utions (Seber 1984). The wing tuft colouration is a discrete character because of the method of scoring (Chapter four), and was often skewed, so to be safe the character was not included in the LDF. The abdominal setal colouration was a simple two-state character, and so could not be distributed normally, this character was also not included in case it affected the performance of the resultant LDFs.

An alternative approach, and the one which was adopted in this analysis, is to manipulate the prior probabilities of species membership according to the fly's score for wing tuft colouration and/or abdominal setal colouration. The method used for deriving the prior probabilities was:

1. The species in a particular analysis were tested for equality of wing tuft colouration and abdominal setal colouration using a non-parametric one-way analysis of variance (Wilcoxon two-sample rank sum test for two species, Kruskal-Wallis test for more than two, using SAS PROC NPAR1WAY). If either of the null hypotheses were rejected then:

2. For the particular analysis, LDFs were calculated using abdominal setal colouration and/or wing tuft colouration. An artificial data set was created of the ten (wing tuft colouration and abdominal setal colouration), five (wing tuft colouration), or two (abdominal setal colouration) possible outcomes of the two characters (Chapter four); this artificial data set was then classified using the LDFs. The resultant posterior probabilities of species membership were then used as prior probabilities of species membership for a fly with that particular combination of characters.

The prior probabilities affected allocation in the following way. Mahalanobis' distance is calculated to each of the reference species and substituted in the following equation:

$$\pi_i \exp(-\frac{1}{2} D_i^2) / \sum \pi_i \exp(-\frac{1}{2} D_i^2)$$

where π_i is the prior probability of belonging to the i -th species ($i=1\dots g$) and D_i^2 is the distance of the fly from the i -th species. This quantity is the posterior probability of belonging to the i -th species, and the g posterior probabilities sum to one. The fly is allocated to the species for which its posterior probability of membership is the largest.

7.2.4 ALLOCATION SCHEMES The second objection to including a non-normal character in a LDF concerns the method used for allocation of unknown flies. The usual method for allocation involves calculating the fly's score on each LDF and allocating it to the species on which its score is highest (Seber 1984). This is equivalent to assigning it to the species to which it has the smallest Mahalanobis' distance (D_i^2), and also equivalent to assigning it to the species for which its posterior probability is highest. This is known as forced allocation (Campbell 1984).

An alternative approach is to calculate a fly's typicality probability of species membership, but for this approach to be used, the data need to conform approximately to multivariate normality and equal dispersion; hence the objection to including the non-normal characters in the analysis.

The typicality probability is the probability associated with the observed Mahalanobis' distance of the fly to each of the reference species. There are different ways of calculating typicality probabilities (Campbell 1984, Ambergen and Schaafsma 1984), but in this

analysis it was decided to use the method of Ambergen and Schaafsma (1984) because this method allows the construction of approximate confidence intervals for the observed distance from each reference species.

The confidence intervals were calculated in the following way:

1. An unbiased estimate of Mahalanobis' distance was calculated:

$$(n-g-p-1)n^{-1} D_i^2 - n^{-1} i$$

2. An estimate of the variance of this unbiased distance was then calculated:

$$(n-g-p-3)^{-1} \{2D_i^4 + 4(n-g-1)n^{-1} D_i^2 + 2p(n-g-1)n^{-2} i\}$$

where n is the total sample size on which the dispersion matrix was based, g is the number of reference species, p is the number of characters and n_i is the sample size of the i -th species.

The following rules were adhered to throughout the analyses in this chapter and in Chapter eight for the typicality probability allocation of flies:

1. If all of the confidence intervals straddled $\alpha=0.01$ then the fly was classified as untypical of the range of reference species, clearly 1% of flies can be expected to be unallocated in this way. Other values of α could be chosen depending on the requirements of a specific analysis.

2. If the confidence intervals for the probability of belonging to one species was greater than for all the others, and did not include 0.01, and none of the other confidence intervals included probabilities of species membership covered by this one, then the fly was classified into that species.

3. If two confidence intervals included a common range of probabilities of species membership, then the fly was allocated onto

the overlap of the two species, unless one of the confidence intervals contained 0.01, in which case the fly was classified into the species whose probability did not include 0.01.

4. If more than two of the higher probability confidence intervals included a common range of probabilities, then the fly was regarded as unclassified, overlapping.

The principal advantage that using typicality probability for allocation has over forced allocation is that a fly which is atypical of all the reference species is not forced into the species to which it is closest. Also, a fly which lies on the overlap region of two or more species is identified as such and is not forced into the species to which it is closest. Because an unbiased estimate of Mahalanobis' distance is used in calculating the typicality probabilities, then some of the bias of a finite sample size is corrected for.

In summary, typicality probability allocation yields more information about the affinities of a particular fly to a particular species, without direct reference to the other species. Rather than having a simple allocate / don't allocate rule, this method of allocation allows a more informed and biologically more realistic allocation decision to be made.

7.2.5 OTHER METHODS

Apart from these methods, the other multivariate statistical techniques used in this analysis have been described in previous sections: Principal components analysis (Chapter three), Canonical variates analysis (Chapter six), and cluster analysis (Chapter three).

The importance of size in discrimination was assessed by introducing the scores along the first principal component of the pooled within-species correlation matrix into the model as a covariable in

a multivariate analysis of covariance (MANCOVA). The first principal component of the pooled within-species correlation matrix is usually a size vector in morphometric studies (Blackith and Reyment 1971), meaning that a unit change in one character is accompanied by a unit change of the same sign in all (Rao 1964). The pooled within-species correlation matrix was used rather than the usual (Total) correlation matrix so that between-species variation and within-species variation were not confused. The analysis was performed using a SAS PROC MATRIX routine to calculate the pooled within-species correlation matrix and to run the principal components analysis whilst SAS PROC GLM was used for the multivariate analysis of covariance.

The assumption of equal dispersion was tested using the likelihood ratio test (Seber 1984). It is known (Layard 1974) that this test is very sensitive even to slight non-normality, and a rejection of the null hypothesis is as likely to be due to departure from multivariate normality as it is to departure from equal dispersion. Despite this criticism of the test, it was still used in this analysis as an indicator of the reliability of the sample statistics obtained. If the null hypothesis was rejected, then caution should be used in the interpretation of results using those statistics.

Strictly, if the null hypothesis is rejected, then each species' dispersion matrix should be calculated and a quadratic discriminant function (QDF) used (Lachenbruch 1975, Seber 1984), or typicality probabilities calculated using the individual species' dispersion matrices rather than the pooled within-species dispersion matrix (Ambergen and Schaafsma 1984, Campbell 1984). This option was not investigated in this analysis for the following reasons:

1. The QDF has poor small sample properties (Seber 1984);

2. The QDF is very sensitive to departures from multivariate normality;
3. More parameters need to be estimated (each species' dispersion matrix rather than the pooled dispersion matrix, as well as the mean vectors), with the result that for the same sample size the precision of estimation of these parameters is reduced (Van Ness and Simpson 1976), which has the effect of broadening the confidence intervals for the typicality probability of species membership (Ambergen and Schaafsma 1984);
4. Greater computation is needed to allocate an unknown fly and more statistics need to be stored in computer memory, which both work against the field applicability of the method.

7.3 RESULTS AND DISCUSSION, TOGO AND BENIN

7.3.1 SPECIES PAIR DISCRIMINATION

In order to deal with the situation where typicality probability allocation results in a fly being allocated to the overlap between two species, then re-allocating the fly in an allocation scheme involving just those two species should increase the chance of correct allocation. Also, in some areas it is known *a priori* that only two species will be expected (e.g. River Gban-Houa, Togo, where *S. sanctipauli* 'Djodji' and *S. squamosum* are the only two species likely to be found), so it would be advantageous to calculate the probability of species membership only in relation to these two species.

For these reasons, statistics for species-pair discrimination will be presented first, followed by those for overall discrimination.

7.3.1.1 Discrimination of *Simulium soubrense* 'Beffa' and *S. damnosum*

The squared Mahalanobis' distance between species using the 25 character set was 31.32 (unbiased $D^2=20.74$, $e_{act}=0.0114$) with a single fly misallocated using resubstitution ($e_{res}=0.0127$). A stepwise discriminant analysis generated a nine character subset which gave a Mahalanobis' squared distance of 26.121 (unbiased $D^2=22.73$, $e_{act}=0.0086$) and no change in the resubstituted error rate ($e_{res}=0.0127$).

The dimension reduction procedure described in section 7.2.1 resulted in a six character subset:

[V9,V14,V18,V19,V27,V28]

i.e. two antennal, two wing and two leg characters, with a Mahalanobis' squared distance of 20.11 (unbiased $D^2=18.282$, $e_{act}=0.0163$) and a single misallocated flies using resubstitution

($e_{res}=0.0127$), which was a *S. damnosum* classified as *S. soubrense* 'Beffa'.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in one misallocated fly (1.27%), two atypical flies (2.53%), one overlapping fly (1.27%) and 75 correctly allocated flies (94.94%).

The null hypothesis of equal wing tuft colouration was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test. The prior probabilities of species membership for each wing tuft colouration category were calculated using the method described in section 7.2.3 and are shown as Table 7.1. When the prior probabilities were adjusted the number of flies misallocated remained at one ($e_{res}=0.0127$).

The null hypothesis of equal dispersion was not rejected at $p<0.001$ using the likelihood ratio test, legitimising the pooling of the dispersion matrices.

The standardised canonical variate (Table 7.2) shows that the main character discriminating these species is antennal length 1, with a relationship between femur length 2 and tibia length 2 also having some effect.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.37,0.20,0.45,0.43,0.47,0.47]

and accounted for 66.4% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 5.1175 to 4.8085 showing that size has only a minor influence in the morphometric differentiation between these species.

The mean vectors are shown as Table 7.3, the linear discriminant functions as Table 7.4 and the pooled within-species dispersion matrix as Table 7.5.

To conclude, there is significant multivariate morphometric differentiation between *S. soubrense* 'Beffa' and *S. damnosum*. This differentiation involves mainly shape variation as in the absence of size there is still good discrimination. Garms and Cheke (1985) considered that these two species were distinguishable using thorax antennal ratios, although their histograms show some overlap. Thus, the character set derived in this analysis is an improvement over previous methods. The six character subset can be expected to classify flies to their correct species in nearly 95% of cases when it is known *a priori* that just this species pair is expected.

Adjusting the prior probabilities of species membership according to a fly's wing tuft colouration does not improve allocation rate, so it is not recommended that this should be done, except in cases of doubt.

Table 7.1.

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------------------------|-------------------------|
| | <i>S. soubrense</i> 'Beffa' | <i>S. damnosum</i> s.s. |
| 1 | 0.0106 | 0.9894 |
| 2 | 0.2986 | 0.7014 |
| 3 | 0.9442 | 0.0558 |
| 4 | 0.9985 | 0.0015 |
| 5 | 1.0 | 0.0 |

Table 7.2

Standardised Canonical Variate for *S. soubrense* 'Beffa' and *S. damnosum* s.s.

| Character | Canonical Variate |
|--------------------|-------------------|
| Antennal Length 1 | 1.5086 |
| Antennal Segment 7 | 0.6834 |
| Wing Width 1 | -0.4247 |
| Wing Width 2 | -0.3936 |
| Femur Length 2 | 0.8102 |
| Tibia Length 2 | -0.9750 |

Table 7.3

Mean Vectors for species *S. soubrense* 'Beffa' and *S. damnosum* s.s.

| Character | <i>S. soubrense</i> 'Beffa' | <i>S. damnosum</i> s.s. |
|--------------------|-----------------------------|-------------------------|
| Antennal Length 1 | 294.17916667 | 259.07790698 |
| Antennal Segment 7 | 43.12444444 | 36.62325581 |
| Wing Width 1 | 961.54361111 | 996.17825581 |
| Wing Width 2 | 1341.42277778 | 1419.20348837 |
| Femur Length 2 | 658.46000000 | 653.67348837 |
| Tibia Length 2 | 611.72000000 | 619.80558140 |

Table 7.4

Linear Discriminant functions for species *S. soubrense* 'Beffa' and *S. damnosum* s.s.

| | <i>S. soubrense</i> 'Beffa' | <i>S. damnosum</i> s.s. |
|----------|-----------------------------|-------------------------|
| CONSTANT | -339.40213292 | -305.70354140 |
| V9 | 0.79149552 | 0.48709948 |
| V14 | 2.88058489 | 2.16075641 |
| V18 | 0.26957828 | 0.30973237 |
| V19 | -0.03963553 | -0.01919285 |
| V27 | 0.37263982 | 0.25241827 |
| V28 | -0.21198021 | -0.06491194 |

Table 7.5

Pooled within-species dispersion matrix for species *S. soubrense* 'Beffa' and *S. damnosum* s.s.

| CHARACTER | V9 | V14 | V18 |
|-----------|--------------|-------------|---------------|
| V9 | 186.83945957 | 12.65833021 | 327.11915002 |
| V14 | 12.65833021 | 7.60652510 | 34.33590606 |
| V18 | 327.11915002 | 34.33590606 | 1973.49030357 |
| V19 | 507.64767399 | 41.91742685 | 2759.17544091 |
| V27 | 262.81697161 | 22.33336249 | 1076.58329496 |
| V28 | 264.95975717 | 22.27992751 | 1038.32222621 |

| CHARACTER | V19 | V27 | V28 |
|-----------|---------------|---------------|---------------|
| V9 | 507.64767399 | 262.81697161 | 264.95975717 |
| V14 | 41.91742685 | 22.33336249 | 22.27992751 |
| V18 | 2759.17544091 | 1076.58329496 | 1038.32222621 |
| V19 | 6012.41826947 | 1779.96639190 | 1650.71783588 |
| V27 | 1779.96639190 | 919.26563866 | 830.83295536 |
| V28 | 1650.71783588 | 830.83295536 | 878.65357221 |

7.3.1.2 Discrimination of Simulium soubrense 'Beffa' and S. sanctipauli 'Djodji'

The squared Mahalanobis' distance between species using the 25 character set was 36.42 (unbiased $D^2=20.64$, $e_{act}=0.0116$) with no flies misallocated using resubstitution ($e_{res}=0.0$). A stepwise discriminant analysis produced a nine character subset which gave a Mahalanobis' squared distance of 18.261 (unbiased $D^2=15.22$, $e_{act}=0.0256$) with a single fly misallocated using resubstitution ($e_{res}=0.0161$).

The dimension reduction technique described in section 7.2.1 resulted in a six character subset:

[V6, V10, V17, V20, V27, V29]

i.e. one head, one antennal, two wing and two leg characters, with a Mahalanobis' squared distance of 10.91 (unbiased $D^2=8.92$, $e_{act}=0.068$) and two misallocated flies using resubstitution ($e_{res}=0.0645$), one into each species.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in two misallocated flies (3.23%), no atypical flies, six overlapping flies (9.68%) and 54 correctly allocated flies (87.1%).

The null hypothesis of equal wing tuft colouration was not rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test. The prior probabilities of species membership for each wing tuft colouration category were therefore not calculated.

The null hypothesis of equal dispersion was not rejected at $p<0.001$ using the likelihood ratio test, legitimising the pooling of the dispersion matrices.

The standardised canonical variate (Table 7.6) shows that the primary contrast discriminating these species is the positively loading characters antennal length 2, wing length 3, and basitarsus length 2, against a negatively loading character head width.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.41,0.37,0.39,0.42,0.44,0.43]

and accounted for 80.1% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 2.5402 to 1.2334 showing that size has a significant influence in discriminating between these species.

The mean vectors are shown as Table 7.7, the linear discriminant functions as Table 7.8 and the pooled within-species dispersion matrix as Table 7.9.

To conclude, there is significant multivariate morphometric differentiation between *S. soubrense* 'Beffa' and *S. sanctipauli*

'Djodji'. Chromosomally, these two species are close (Chapter three, Post 1986, Meredith *et al.* 1984), so it is not surprising that they are close together morphometrically. Previous work using morphometrics of the *S. damnosum* complex have not distinguished between these species (e.g. Garms and Cheke 1985), but recent work on the epidemiological importance of *S. sanctipauli* 'Djodji' (Cheke and Denke 1988) shows that identification of *S. sanctipauli* 'Djodji' is important. Much of the separation between these species involves size, so that once a wider range of variation has been examined, the error rate will probably increase.

Table 7.6

Standardised Canonical Variate for *S. soubrense* 'Beffa' and *S. sanctipauli* 'Djodji'

| Character | Canonical Variate |
|-------------------|-------------------|
| Head Width | -1.3623 |
| Antennal Length 2 | 1.0467 |
| Wing Length 2 | -0.6889 |
| Wing Length 3 | 1.1002 |
| Femur Length 2 | -0.5752 |
| Basitarsus Length | 1.8089 |

Table 7.7

Mean Vectors for species *S. soubrense* 'Beffa' and *S. sanctipauli* 'Djodji'

| Character | <i>S. soubrense</i> 'Beffa' | <i>S. sanctipauli</i> |
|---------------------|-----------------------------|-----------------------|
| Head Width | 815.04746667 | 838.51038462 |
| Antennal Length 2 | 442.52666667 | 476.10461538 |
| Wing Length 2 | 438.70000000 | 467.96769231 |
| Wing Length 3 | 1409.63763889 | 1534.97576923 |
| Femur Length 2 | 658.46000000 | 697.88307692 |
| Basitarsus Length 2 | 320.77500000 | 350.32500000 |

Table 7.8

Linear Discriminant functions for species *S. soubrense* 'Beffa' and *S. sanctipauli* 'Djodji'

| Character | <i>S. soubrense</i> 'Beffa' | <i>S. sanctipauli</i> 'Djodji' |
|-----------|-----------------------------|--------------------------------|
| CONSTANT | -287.63555221 | -331.59567479 |
| V6 | 0.19155864 | 0.08060720 |
| V10 | 0.66202707 | 0.79023419 |
| V17 | -0.13911607 | -0.21223907 |
| V20 | 0.03484206 | 0.07186381 |
| V27 | 0.03900777 | -0.01079215 |
| V29 | 0.35042500 | 0.61631544 |

Table 7.9

Pooled within-species dispersion matrix for species *S. soubrense* 'Beffa' and *S. sanctipauli*

| Character | V6 | V10 | V17 |
|-----------|---------------|--------------|---------------|
| V6 | 1408.89212946 | 532.60455016 | 649.46063498 |
| V10 | 532.60455016 | 400.41592410 | 327.60362462 |
| V17 | 649.46063498 | 327.60362462 | 695.39090769 |
| V20 | 2163.13096102 | 908.01100124 | 1416.44027244 |
| V27 | 989.61622415 | 454.97687385 | 656.58308308 |
| V29 | 482.10925950 | 204.33990000 | 340.11345000 |

| Character | V20 | V27 | V29 |
|-----------|---------------|---------------|--------------|
| V6 | 2163.13096102 | 989.61622415 | 482.10925950 |
| V10 | 908.01100124 | 454.97687385 | 204.33990000 |
| V17 | 1416.44027244 | 656.58308308 | 340.11345000 |
| V20 | 5112.05735181 | 1891.56033064 | 968.23234375 |
| V27 | 1891.56033064 | 978.09372923 | 442.22190000 |
| V29 | 968.23234375 | 442.22190000 | 255.33056250 |

7.3.1.3 Discrimination of Simulium soubrense 'Beffa' and S. squamosum

The squared Mahalanobis' distance between species using the 25 character set was 29.21 (unbiased $D^2=22.83$, $e_{act}=0.0073$) with a single fly misallocated using resubstitution ($e_{res}=0.0083$). A stepwise discriminant analysis produced a 12 character subset which gave a Mahalanobis' squared distance of 27.38 (unbiased $D^2=24.39$, $e_{act}=0.0068$) with one fly misallocated using resubstitution ($e_{res}=0.0083$).

The dimension reduction technique described in section 7.2.1 resulted in a four character subset:

[V6,V10,V18,V20]

i.e. one head, one antennal and two wing characters, with a Mahalanobis' squared distance of 19.0 (unbiased $D^2=18.2$, $e_{act}=0.00165$) and one misallocated fly using resubstitution

($e_{res}=0.0083$), which was a *S. squamosum* misallocated into *S. soubrense* 'Beffa'

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in one misallocated fly (0.83%), two atypical flies (1.65%) no overlapping flies and 118 correctly allocated flies (97.52%).

The null hypothesis of equal wing tuft colouration was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test. The prior probabilities of species membership for each wing tuft colouration category were calculated using the method described in section 7.2.3 and are shown as Table 7.10. When the prior probabilities were adjusted according to wing tuft colouration, then three flies were misallocated using resubstitution ($e_{res}=0.0248$).

The null hypothesis of equal dispersion was rejected at $p<0.001$ using the likelihood ratio test, however the pooled dispersion matrix was still used for the practical and statistical reasons given in section 7.2.5.

The standardised canonical variate (Table 7.11) shows that the main contrast discriminating these species is the positively loading wing characters wing width 1 and wing length 3, and the negatively loading character antennal length 2.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.49,0.48,0.51,0.52]

and accounted for 84.4% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 4.0397 to 2.4333 showing that

size has an important influence in discriminating between these species, but that shape differences are as significant.

The mean vectors are shown as Table 7.12, the linear discriminant functions as Table 7.13 and the pooled within-species dispersion matrix as Table 7.14.

To conclude, there is significant multivariate morphometric discrimination between *S. soubrense* 'Beffa' and *S. squamosum*. This differentiation involves a combination of size and shape variation. Garms and Cheke (1985) consider this species pair to be very difficult to distinguish using thorax antennal ratios, so the character subset derived in this analysis is a considerable improvement. The four character subset can be expected to classify flies to their correct species in over 97% of cases when it is known *a priori* that just this species pair is expected.

Adjusting the prior probabilities of species membership according to a fly's wing tuft colouration does not improve allocation rate, due to *S. soubrense* 'Beffa' having wing tufts falling into all five colour categories (Chapter four). Thus, *S. soubrense* 'Beffa' with colour categories one or two are strongly penalised against 'own-group' membership. Such flies were found in this data set, and have been reported in previous work (Meredith *et al.* 1984) Therefore it is not recommended that the prior probabilities should be altered.

Table 7.10

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------------------------|---------------------|
| | <i>S. soubrense</i> 'Beffa' | <i>S. squamosum</i> |
| 1 | 0.0062 | 0.9938 |
| 2 | 0.2513 | 0.7487 |
| 3 | 0.9473 | 0.0527 |
| 4 | 0.9990 | 0.0010 |
| 5 | 1.0000 | 0.0000 |

Table 7.11

Standardised Canonical Variate for *S. soubrense* 'Beffa' and *S. squamosum*

| Character | Canonical Variate |
|-------------------|-------------------|
| Head Width | -0.4915 |
| Antennal Length 2 | -1.3310 |
| Wing Width 1 | 1.3799 |
| Wing Length 3 | 1.1973 |

Table 7.12

Mean Vectors for species *S. soubrense* 'Beffa' and *S. squamosum*

| Character | <i>S. soubrense</i> 'Beffa' | <i>S. squamosum</i> |
|-------------------|-----------------------------|---------------------|
| Head Width | 815.04746667 | 839.65947529 |
| Antennal Length 2 | 442.52666667 | 425.89835294 |
| Wing Width 1 | 961.54361111 | 1087.65100000 |
| Wing Length 3 | 1409.63763889 | 1574.30000000 |

Table 7.13

Linear Discriminant functions for species *S. soubrense* 'Beffa' and *S. squamosum*

| Character | <i>S. soubrense</i> 'Beffa' | <i>S. squamosum</i> |
|-----------|-----------------------------|---------------------|
| CONSTANT | -190.90523318 | -194.88882867 |
| V6 | 0.28684170 | 0.23993834 |
| V10 | 0.43286097 | 0.20806258 |
| V18 | -0.02160380 | 0.04978747 |
| V20 | -0.01614504 | 0.02893111 |

Table 7.14

Pooled within-species dispersion matrix for species *S. soubrense* 'Beffa' and *S. squamosum*

| Character | V6 | V10 | V18 | V20 |
|-----------|---------------|---------------|---------------|---------------|
| V6 | 1975.76775752 | 799.57472278 | 2167.22308444 | 3156.97831053 |
| V10 | 799.57472278 | 613.25246529 | 1153.60869894 | 1694.35477955 |
| V18 | 2167.22308444 | 1153.60869894 | 3782.23936992 | 4815.58203830 |
| V20 | 3156.97831053 | 1694.35477955 | 4815.58203830 | 7760.93115189 |

7.3.1.4 Discrimination of *Simulium damnosum* and *S. sanctipauli* 'Djodji'

The squared Mahalanobis' distance between species using the 25 character set was 52.42 (unbiased $D^2=32.08$, $e_{act}=0.0023$) with a single fly misallocated using resubstitution ($e_{res}=0.015$). A stepwise discriminant analysis produced a 7 character subset which gave a Mahalanobis' squared distance of 43.15 (unbiased $D^2=37.996$, $e_{act}=0.001$) with no flies misallocated using resubstitution ($e_{res}=0.0$).

The dimension reduction technique described in section 7.2.1 resulted in a four character subset:

[V6, V9, V13, V29]

i.e. one head, two antennal and one leg character, with a Mahalanobis' squared distance of 36.4 (unbiased $D^2=33.68$, $e_{act}=0.00186$) and no misallocated flies using resubstitution ($e_{res}=0.0$).

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in no misallocated flies, one atypical fly (1.45%), no overlapping flies and 68 correctly allocated flies (98.55%).

The null hypothesis of equal wing tuft colouration was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test. The prior probabilities of species membership for each wing tuft colouration category were calculated using the method described in section 7.2.3 and are shown as Table 7.15. When the prior probabilities were adjusted according to wing tuft colouration, no flies were misallocated using resubstitution ($e_{res}=0.0$).

The null hypothesis of equal dispersion was not rejected at $p<0.001$ using the likelihood ratio test, legitimising the use of the pooled dispersion matrix.

The standardised canonical variate (Table 7.16) shows that three of the characters load positively (antennal length 1, antennal segment 6, basitarsus length 2) and one negatively (head width). Antennal length relative to other characters is the most important discriminatory character.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

$$[0.56, 0.52, 0.36, 0.54]$$

and accounted for 58.6% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 8.803 to 3.351 showing that

size has a significant influence in discriminating between these species, but that shape differences are still very important in the absence of size variation.

The mean vectors are shown as Table 7.17, the linear discriminant functions as Table 7.18 and the pooled within-species dispersion matrix as Table 7.19.

To conclude, there is significant multivariate morphometric differentiation between *S. damnosum* and *S. sanctipauli* 'Djodji'. This differentiation involves a combination of size and shape variation, but shape is very important as a discriminatory character. The four character subset can be expected to classify flies to their correct species in nearly 99% of cases when it is known *a priori* that just this species pair is expected.

Adjusting the prior probabilities of species membership according to a fly's wing tuft colouration does not improve allocation rate, so it is recommended that this should be done only in cases of doubt following typicality probability allocation.

Table 7.15

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-------------------------|--------------------------------|
| | <i>S. damnosum</i> s.s. | <i>S. sanctipauli</i> 'Djodji' |
| 1 | 1.0000 | 0.0000 |
| 2 | 0.9738 | 0.0262 |
| 3 | 0.0138 | 0.9862 |
| 4 | 0.0000 | 1.0000 |
| 5 | 0.0000 | 1.0000 |

Table 7.16

Standardised Canonical Variate for *S. damnosum s.s.* and *S. sanctipauli* 'Djodji'

| Character | Canonical Variate |
|--------------------|-------------------|
| Head Width | -0.7509 |
| Antennal Length 1 | 1.8187 |
| Antennal Segment 6 | 1.2273 |
| Basitarsus Length | 0.5989 |

Table 7.17

Mean Vectors for species *S. damnosum s.s.* and *S. sanctipauli* 'Djodji'

| | <i>S. damnosum s.s.</i> | <i>S. sanctipauli</i> |
|---------------------|-------------------------|-----------------------|
| Head Width | 812.94496744 | 838.51038462 |
| Antennal Length 1 | 259.07790698 | 317.62500000 |
| Antennal Segment 6 | 37.05581395 | 49.64769231 |
| Basitarsus Length 2 | 320.16279070 | 350.32500000 |

Table 7.18

Linear Discriminant functions for species *S. damnosum s.s.* and *S. sanctipauli* 'Djodji'

| | <i>S. damnosum s.s.</i> | <i>S. sanctipauli</i> 'Djodji' |
|----------|-------------------------|-----------------------------------|
| CONSTANT | -376.76328848 | -479.42883704 |
| V6 | 0.35764041 | 0.22851342 |
| V9 | 0.65331761 | 1.00564940 |
| V13 | 1.03505757 | 2.11469538 |
| V29 | 0.79699938 | 0.97862744 |

Table 7.19

Pooled within-species dispersion matrix for species *S. damnosum* s.s. and *S. sanctipauli* 'Djodji'

| Character | V6 | V9 | V13 | V29 |
|-----------|---------------|--------------|-------------|--------------|
| V6 | 1091.36893129 | 211.47417506 | 32.18249829 | 315.12796654 |
| V9 | 211.47417506 | 155.38130988 | 14.73961229 | 83.65941644 |
| V13 | 32.18249829 | 14.73961229 | 9.39496878 | 7.76922839 |
| V29 | 315.12796654 | 83.65941644 | 7.76922839 | 181.63442858 |

7.3.1.5 Discrimination of *Simulium damnosum* and *S. squamosum*

The squared Mahalanobis' distance between species using the 25 character set was 16.96 (unbiased $D^2=13.46$, $e_{act}=0.0331$) with two flies misallocated using resubstitution ($e_{res}=0.01563$). A stepwise discriminant analysis produced an 11 character subset which gave a Mahalanobis' squared distance of 15.304 (unbiased $D^2=13.85$, $e_{act}=0.03139$) with two flies misallocated using resubstitution ($e_{res}=0.01563$).

The dimension reduction technique described in section 7.2.2 resulted in a nine character subset:

[V3,V4,V9,V12,V17,V19,V23,V28,V29]

i.e. two thorax, two antennal, two wing and three leg characters, with a Mahalanobis' squared distance of 13.32 (unbiased $D^2=12.26$, $e_{act}=0.04$) and three misallocated flies using resubstitution ($e_{res}=0.0234$), two misallocated into *S. squamosum*, the other into *S. damnosum*.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in two misallocated flies (1.56%), three atypical flies (2.34%) six overlapping flies (4.69%) and 117 correctly allocated flies (91.41%).

The null hypothesis of equal wing tuft colouration was not rejected at $p < 0.001$ using a Wilcoxon two-sample rank sum test. The prior probabilities of species membership for each wing tuft colouration category were therefore not calculated.

The null hypothesis of equal dispersion was not rejected at $p < 0.001$ using the likelihood ratio test, legitimising the use of the pooled dispersion matrix.

The standardised canonical variate (Table 7.20) shows that the characters of most importance in discrimination are the positively loading characters: tibia length 1, basitarsus length 2, and the negatively loading characters: thorax width and tibia length 2.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.33, 0.36, 0.28, 0.20, 0.34, 0.34, 0.37, 0.37, 0.36]

and accounted for 73.5% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 3.0186 to 2.0499 showing that size has a significant influence in discriminating between these species, but that shape differences are more important.

The mean vectors are shown as Table 7.21, the linear discriminant functions as Table 7.22 and the pooled within-species dispersion matrix as Table 7.23.

To conclude, there is significant multivariate morphometric differentiation between *S. damnosum* and *S. squamosum*. This differentiation involves a combination of size and shape variation, but shape is very important as a discriminatory character. In previous work (Garms and Cheke 1985) these species were considered to be very difficult to distinguish using thorax antennal ratios, so that the

character subset derived in this analysis is an improvement. However, nine characters is quite large, so it might be preferable to rely on the overall discrimination statistics described in section 7.3.2 for identification of this species pair.

Table 7.20

Standardised Canonical Variate for *S. damnosum s.s.* and *S. squamosum*

| Character | Canonical Variate |
|--------------------|-------------------|
| Thorax Length | 0.4729 |
| Thorax Width | -1.6522 |
| Antennal Length 1 | 0.2601 |
| Antennal Segment 5 | 0.4221 |
| Wing Length 2 | -0.3824 |
| Wing Width 2 | 0.4267 |
| Tibia Length 1 | 1.1911 |
| Tibia Length 2 | -1.3827 |
| Basitarsus Length | 1.9785 |

Table 7.21

Mean Vectors for species *S. damnosum s.s.* and *S. squamosum*

| | <i>S. damnosum s.s.</i> | <i>S. squamosum</i> |
|---------------------|-------------------------|---------------------|
| Thorax Length | 631.52014884 | 680.26941882 |
| Thorax Width | 907.70187907 | 927.77268706 |
| Antennal Length 1 | 259.07790698 | 282.91058824 |
| Antennal Segment 5 | 34.74883721 | 39.35905882 |
| Wing Length 2 | 454.47069767 | 476.02447059 |
| Wing Width 2 | 1419.20348837 | 1522.57300000 |
| Tibia Length 1 | 689.37209302 | 749.57647059 |
| Tibia Length 2 | 619.80558140 | 662.40564706 |
| Basitarsus Length 2 | 320.16279070 | 357.69882353 |

Table 7.22

Linear Discriminant functions for species *S. damnosum s.s.* and *S. squamosum*

| | <i>S. damnosum s.s.</i> | <i>S. squamosum</i> |
|----------|-------------------------|---------------------|
| CONSTANT | -238.68582302 | -270.07005463 |
| V3 | -0.15676510 | -0.12640568 |
| V4 | 0.17864883 | 0.07054880 |
| V9 | 0.90987717 | 0.95944660 |
| V12 | 0.41001607 | 0.82246741 |
| V17 | 0.03658488 | -0.00681295 |
| V19 | 0.26517257 | 0.28239191 |
| V23 | -0.21673943 | -0.13174467 |
| V28 | 0.02150117 | -0.09219862 |
| V29 | -0.28934526 | -0.02802901 |

Table 7.23

Pooled within-species dispersion matrix for species *S. damnosum* s.s. and *S. squamosum*

| Character | V3 | V4 | V9 |
|-----------|---------------|---------------|---------------|
| V3 | 2719.42384454 | 2343.02705442 | 463.99135654 |
| V4 | 2343.02705442 | 3045.02475869 | 497.58785243 |
| V9 | 463.99135654 | 497.58785243 | 240.87287287 |
| V12 | 58.31889137 | 74.78838804 | 19.86120827 |
| V17 | 1161.84786378 | 1311.54845189 | 259.72038999 |
| V19 | 2760.74649043 | 3398.17483861 | 567.86963583 |
| V23 | 1741.21253327 | 2092.51304810 | 427.40714813 |
| V28 | 1664.75317336 | 1917.48843647 | 405.64447794 |
| V29 | 828.83677365 | 999.41286012 | 209.29048183 |
| Character | V12 | V17 | V19 |
| V3 | 58.31889137 | 1161.84786378 | 2760.74649043 |
| V4 | 74.78838804 | 1311.54845189 | 3398.17483861 |
| V9 | 19.86120827 | 259.72038999 | 567.86963583 |
| V12 | 9.24125212 | 26.34928804 | 99.18742789 |
| V17 | 26.34928804 | 937.10625857 | 1815.13373298 |
| V19 | 99.18742789 | 1815.13373298 | 5824.41303244 |
| V23 | 57.79532002 | 1065.76385538 | 2769.46224275 |
| V28 | 51.22733913 | 994.32845783 | 2448.53392002 |
| V29 | 24.78239719 | 521.21037749 | 1298.53001096 |
| Character | V23 | V28 | V29 |
| V3 | 1741.21253327 | 1664.75317336 | 828.83677365 |
| V4 | 2092.51304810 | 1917.48843647 | 999.41286012 |
| V9 | 427.40714813 | 405.64447794 | 209.29048183 |
| V12 | 57.79532002 | 51.22733913 | 24.78239719 |
| V17 | 1065.76385538 | 994.32845783 | 521.21037749 |
| V19 | 2769.46224275 | 2448.53392002 | 1298.53001096 |
| V23 | 1815.06318296 | 1606.66280946 | 824.62797616 |
| V28 | 1606.66280946 | 1574.31333928 | 771.22664678 |
| V29 | 824.62797616 | 771.22664678 | 450.30670911 |

7.3.1.6 Discrimination of Simulium sanctipauli 'Djodji' and S. squamosum

The squared Mahalanobis' distance between species using the 25 character set was 31.1 (unbiased $D^2=23.68$, $e_{act}=0.0075$) with no flies misallocated using resubstitution ($e_{res}=0.0$). A stepwise discriminant analysis produced a seven character subset which gave a Mahalanobis' squared distance of 24.66 (unbiased $D^2=22.85$,

$e_{act}=0.00842$) with two flies misallocated using resubstitution ($e_{res}=0.018$).

The dimension reduction technique described in section 7.2.2 resulted in a three character subset:

[V10,V18,V22]

i.e. one antennal, one wing and one leg character, with a Mahalanobis' squared distance of 15.5 (unbiased $D^2=15.12$, $e_{act}=0.00259$) and two misallocated flies using resubstitution ($e_{res}=0.018$), both misallocated into *S. sanctipauli* 'Djodji'.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in two misallocated flies (1.8%), no atypical flies, no overlapping flies and 109 correctly allocated flies (98.2%).

The null hypothesis of equal wing tuft colouration was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test. The prior probabilities of species membership for each wing tuft colouration category were calculated and are shown as Table 7.24. When the prior probabilities of species membership were adjusted for each wing tuft colouration category, the resubstituted error rate fell to one fly misallocated into *S. sanctipauli* 'Djodji' ($e_{res}=0.009$).

The null hypothesis of equal dispersion was not rejected at $p<0.001$ using the likelihood ratio test, legitimising the use of the pooled dispersion matrix.

The standardised canonical variate (Table 7.25) shows that the most important character in discrimination between these species is antennal length 2, which contrasts with the other two characters.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.55,0.58,0.59]

and accounted for 85.6% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 2.8403 to 2.8257 showing that size has negligible influence on discriminating between these species.

The mean vectors are shown as Table 7.26, the linear discriminant functions as Table 7.27 and the pooled within-species dispersion matrix as Table 7.28.

To conclude, there is significant multivariate morphometric differentiation between *S. sanctipauli* 'Djodji' and *S. squamosum*. The excellent discrimination between these species is encouraging as *S. sanctipauli* 'Djodji' has often been found in sympatry with *S. squamosum* (Table 2.1). Cheke and Denke (1988) found that *S. sanctipauli* 'Djodji' was a more efficient vector than *S. squamosum* so that successful identification of these two species is important. The four character subset correctly identifies over 98% of flies and is not influenced by size variation.

Adjusting the prior probabilities only slightly improves error rate, so it is recommended that this be done only in cases of doubt following typicality probability allocation.

Table 7.24

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------------------------------|---------------------|
| | <i>S. sanctipauli</i> 'Djodji' | <i>S. squamosum</i> |
| 1 | 0.0000 | 1.0000 |
| 2 | 0.0281 | 0.9719 |
| 3 | 0.9716 | 0.0284 |
| 4 | 1.0000 | 0.0000 |
| 5 | 1.0000 | 0.0000 |

Table 7.25

Standardised Canonical Variate for *S. sanctipauli* 'Djodji' and *S. squamosum*

| Character | Canonical Variate |
|-------------------|-------------------|
| Antennal Length 2 | -2.0029 |
| Wing Width 1 | 0.7875 |
| Femur Length 1 | 0.6119 |

Table 7.26

Mean Vectors for species *S. sanctipauli* 'Djodji' and *S. squamosum*

| | <i>S. sanctipauli</i> 'Djodji' | <i>S. squamosum</i> |
|-------------------|-----------------------------------|---------------------|
| Antennal Length 2 | 476.10461538 | 425.89835294 |
| Wing Width 1 | 1041.22711538 | 1087.65100000 |
| Femur Length 1 | 660.41538462 | 678.32329412 |

Table 7.27

Linear Discriminant functions for species *S. sanctipauli* 'Djodji' and *S. squamosum*

| | <i>S. sanctipauli</i> 'Djodji' | <i>S. squamosum</i> |
|----------|-----------------------------------|---------------------|
| CONSTANT | -208.20478421 | -189.45004107 |
| V10 | 0.66157179 | 0.41817336 |
| V18 | 0.21633488 | 0.26616874 |
| V22 | -0.18749041 | -0.13076059 |

Table 7.28

Pooled within-species dispersion matrix for species *S. sanctipauli* 'Djodji' and *S. squamosum*

| Character | V10 | V18 | V22 |
|-----------|---------------|---------------|---------------|
| V10 | 602.16406987 | 1039.78894941 | 785.17486436 |
| V18 | 1039.78894941 | 3525.12490412 | 2182.91364838 |
| V22 | 785.17486436 | 2182.91364838 | 1766.88772958 |

7.3.2 OVERALL DISCRIMINATION The matrix of Mahalanobis' squared distances between species using the full 25 character set excluding wing tuft colouration, abdominal setal colouration and basitarsal spine number is shown as Table 7.29. All of these distances are significant at $p < 0.001$. Examining this matrix shows that the two members of the *S. sanctipauli* subcomplex are closer to each other than either is to either *S. damnosum* or *S. squamosum*. The table of re-substitutions is given in Table 7.30. The data set was too large to obtain an estimate of error rate using the 'leave-one-out' method. Because of this, the resubstituted error rate, $e_{res} = 0.0263$ might be optimistic.

A stepwise discriminant analysis of the 25 character set resulted in an initial subset of 15 characters. The method described in section 7.2.3 was then used to reduce the number of characters from the 15 character subset to a nine character subset:

[V4, V6, V9, V18, V20, V23, V27, V28, V29]

i.e. one thorax, one head, one antennal, two wing and four leg characters. The matrix of Mahalanobis' squared distances resulting from this character subset is shown as Table 7.31. The same pattern of species relationships holds in this lower dimensional space as in the 25-character space, with *S. sanctipauli* 'Djodji' and *S. soubrense* 'Beffa' still closer to each other than either is to the other species.

Table 7.32 gives the table of reclassifications using resubstitution and the 'leave-one-out' method of error rate estimation. The resubstituted error rate using the nine character subset was 0.0842, higher than for the full character set, the estimate of error rate

using the 'leave-one-out' method was 0.1105, showing that the resubstituted error rate is quite biased.

Allocation using the typicality probability method described in section 7.2.4 resulted in 154 (81.05%) correctly allocated flies, 10 (5.26%) misidentified flies, 20 (10.53%) flies lying on an overlap region and 6 (3.16%) flies being untypical of any of the reference species, with atypicality defined at $\alpha=0.01$. The larger than expected number of atypical flies was most likely due to the fact that *S. squamosum* showed considerable variation in size (see Chapter six), as five of the atypical flies were *S. squamosum*. Of the twenty flies which were overlapping 8 remained in an overlap region and 12 were correctly allocated when they were allocated using the relevant species-pair statistics described in section 7.3.1, bringing the number of flies correctly allocated up to 166 (87.37%).

The null hypothesis of equal wing tuft colouration was rejected using a Kruskal-Wallis test at $p<0.001$, the null hypothesis of equal abdominal setal colouration was not tested as all flies in these samples were character state one for this character (Chapter four). The prior probabilities of species membership according to a fly's wing tuft colouration were therefore calculated and are shown as Table 7.33. When the prior probabilities were adjusted in the way described in section 7.2.3, then 14 flies were classified incorrectly ($e_{res}=0.0737$).

The null hypothesis of equal dispersion was not rejected at $p<0.001$ using the likelihood ratio test, legitimising the use of the pooled within-species dispersion matrix.

The standardised canonical variates (Table 7.34, with the species' means along each canonical variate shown as Table 7.35) shows the

characters of importance in discriminating between these species. The first canonical variate with a canonical root of 2.91 is dominated by antennal length 1, with most of the other characters having small negative loadings. This vector discriminates predominately the species pair *S. soubrense* 'Beffa' / *S. sanctipauli* 'Djodji' from the species pair *S. damnosum* / *S. squamosum*. The second canonical variate is a more complex vector but is basically a contrast between the two positively loading characters tibia length 1 and basitarsus length 2 and the two negatively loading characters thorax width and tibia length 2. The canonical root associated with this canonical variate was 2.173, which is high relative to the first canonical root (together they account for nearly 95% of total variance). This canonical variate discriminates mainly *S. damnosum* from the other species. The final canonical variate, with a canonical root of only 0.279, whilst being statistically significant is probably of no biological importance (Campbell 1982), because no single species is discriminated along its length.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.33,0.33,0.26,0.33,0.34,0.35,0.35,0.35,0.34]

and accounted for 83% of pooled within-species variation. When the scores computed along this vector were introduced into the model as a covariable the canonical roots fell from (2.9077,2.1729,0.2786) to (2.8209,1.2079,0.2279). Thus it is clear that the first canonical variate is very little influenced by size variation, whereas the second canonical variate is influenced by size variation to a considerable extent, although there is still significant discrimination along this vector in the absence of size.

The mean vectors are shown as Table 7.36, the pooled within-species dispersion matrix as Table 7.37, and the linear discriminant functions as Table 7.38.

Table 7.29

Matrix of Mahalanobis' distances between species, 25 character set

| | <i>S. soubrense</i> | <i>S. damnosum s.s</i> | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
|------------------------|---------------------|------------------------|-----------------------|---------------------|
| <i>S. soubrense</i> | 0.0 | | | |
| <i>S. damnosum s.s</i> | 27.74 | 0.0 | | |
| <i>S. sanctipauli</i> | 11.88 | 50.11 | 0.0 | |
| <i>S. squamosum</i> | 22.88 | 19.06 | 27.45 | 0.0 |

Table 7.30

Table of re-classifications, using resubstitution, 25 character set

| | <i>S. soubrense</i> | <i>S. damnosum s.s</i> | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
|------------------------|---------------------|------------------------|-----------------------|---------------------|
| <i>S. soubrense</i> | 35 | 0 | 1 | 0 |
| <i>S. damnosum s.s</i> | 1 | 41 | 0 | 1 |
| <i>S. sanctipauli</i> | 0 | 0 | 26 | 0 |
| <i>S. squamosum</i> | 0 | 1 | 1 | 83 |

$$e_{res} = 5/190 = 0.0263$$

Table 7.31

Matrix of Mahalanobis' distances between species, 9 character subset

| | <i>S. soubrense</i> | <i>S. damnosum s.s</i> | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
|------------------------|---------------------|------------------------|-----------------------|---------------------|
| <i>S. soubrense</i> | 0.0 | | | |
| <i>S. damnosum s.s</i> | 16.03 | 0.0 | | |
| <i>S. sanctipauli</i> | 6.31 | 25.43 | 0.0 | |
| <i>S. squamosum</i> | 16.66 | 12.64 | 14.95 | 0.0 |

Table 7.32

Table of re-classifications, using resubstitution (and 'leave-one-out'), 9 character subset

| | <i>S. soubrense</i> | <i>S. damnosum s.s</i> | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
|------------------------|---------------------|------------------------|-----------------------|---------------------|
| <i>S. soubrense</i> | 31(30) | 0(0) | 5(6) | 0(0) |
| <i>S. damnosum s.s</i> | 2(2) | 39(38) | 0(0) | 2(3) |
| <i>S. sanctipauli</i> | 0(2) | 0(0) | 26(24) | 0(0) |
| <i>S. squamosum</i> | 2(2) | 3(4) | 2(4) | 78(75) |

$$e_{\text{res}} = 16/190 = 0.0842$$

$$e_c = 21/190 = 0.1105$$

Table 7.33

Prior probability of species membership for each wing tuft colour category

| Wing Tuft Colour | <i>S. soubrense</i> | <i>S. damnosum s.s</i> | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
|------------------|---------------------|------------------------|-----------------------|---------------------|
| 1 | 0.014 | 0.5102 | 0.001 | 0.4882 |
| 2 | 0.1293 | 0.3904 | 0.0274 | 0.4530 |
| 3 | 0.6038 | 0.0155 | 0.3588 | 0.0218 |
| 4 | 0.3747 | 0.0001 | 0.6251 | 0.0001 |
| 5 | 0.176 | 0.0 | 0.824 | 0.0 |

Table 7.34

Standardised Canonical Variates

| Character | CV I | CV II | CV III |
|-------------------|---------|---------|---------|
| Thorax Width | -0.1382 | -1.1370 | 0.9800 |
| Head Width | 0.3349 | -0.4896 | -0.4974 |
| Antennal Length 1 | 1.8284 | 0.5821 | 0.5528 |
| Wing Width 1 | -0.8490 | 0.6044 | 0.2498 |
| Wing Length 3 | -0.4708 | 0.2720 | 1.2369 |
| Tibia Length 1 | -0.6883 | 1.0006 | -1.8473 |
| Femur Length 2 | 0.2425 | 0.5240 | -2.5364 |
| Tibia Length 2 | 0.1017 | -1.6807 | 2.6348 |
| Basitarsus Length | 0.0327 | 1.6664 | -0.0822 |

Table 7.35

Species Means on Canonical Variates

| Species | CV I | CV II | CV III |
|-------------------------|---------|---------|---------|
| <i>S. soubrense</i> | 2.3471 | -0.7959 | -0.7467 |
| <i>S. damnosum s.s.</i> | -1.1964 | -2.3222 | 0.3222 |
| <i>S. sanctipauli</i> | 2.5281 | 1.0122 | 0.9882 |
| <i>S. squamosum</i> | -1.1621 | 1.2022 | -0.1490 |

Table 7.36

Mean Vectors

| Character | <i>S. soubrense</i> | <i>S. damnosum s.s.</i> | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
|-------------------|---------------------|-------------------------|-----------------------|---------------------|
| Thorax Width | 872.15600000 | 907.70187907 | 914.47961538 | 927.77268706 |
| Head Width | 815.04746667 | 812.94496744 | 838.51038462 | 839.65947529 |
| Antennal Length 1 | 294.17916667 | 259.07790698 | 317.62500000 | 282.91058824 |
| Wing Width 1 | 961.54361111 | 996.17825581 | 1041.22711538 | 1087.65100000 |
| Wing Length 3 | 1409.6376388 | 1452.37523256 | 1534.97576923 | 1574.30000000 |
| Tibia Length 1 | 685.52000000 | 689.37209302 | 730.24153846 | 749.57647059 |
| Femur Length 2 | 658.46000000 | 653.67348837 | 697.88307692 | 708.19058824 |
| Tibia Length 2 | 611.72000000 | 619.80558140 | 659.46923077 | 662.40564706 |
| Basitarsus Length | 320.77500000 | 320.16279070 | 350.32500000 | 357.69882353 |

Table 7.37

Pooled within-species dispersion matrix

| Character | V4 | V6 | V9 |
|-----------|---------------|---------------|---------------|
| V4 | 2703.98603061 | 1785.83050939 | 443.98971197 |
| V6 | 1785.83050939 | 1657.19398678 | 356.99024657 |
| V9 | 443.98971197 | 356.99024657 | 218.59357585 |
| V18 | 2193.77677851 | 1641.83501162 | 459.42494772 |
| V20 | 3177.78803740 | 2437.04611390 | 692.30285860 |
| V23 | 1790.24030011 | 1353.68431045 | 385.62531003 |
| V27 | 1686.60616466 | 1291.23179439 | 370.05397465 |
| V28 | 1620.27957036 | 1233.27917780 | 362.93150118 |
| V29 | 849.82461540 | 643.94857901 | 181.67222156 |
| Character | V18 | V20 | V23 |
| V4 | 2193.77677851 | 3177.78803740 | 1790.24030011 |
| V6 | 1641.83501162 | 2437.04611390 | 1353.68431045 |
| V9 | 459.42494772 | 692.30285860 | 385.62531003 |
| V18 | 2882.78154798 | 3578.93882920 | 1801.24154958 |
| V20 | 3578.93882920 | 5995.15954680 | 2675.27774004 |
| V23 | 1801.24154958 | 2675.27774004 | 1564.32903759 |
| V27 | 1714.13777818 | 2520.76676937 | 1379.39450559 |
| V28 | 1667.97585554 | 2486.64200623 | 1361.36720442 |
| V29 | 890.86306685 | 1316.36304299 | 699.11470965 |
| Character | V27 | V28 | V29 |
| V4 | 1686.60616466 | 1620.27957036 | 849.82461540 |
| V6 | 1291.23179439 | 1233.27917780 | 643.94857901 |
| V9 | 370.05397465 | 362.93150118 | 181.67222156 |
| V18 | 1714.13777818 | 1667.97585554 | 890.86306685 |
| V20 | 2520.76676937 | 2486.64200623 | 1316.36304299 |
| V23 | 1379.39450559 | 1361.36720442 | 699.11470965 |
| V27 | 1402.44246022 | 1303.00588678 | 668.48042280 |
| V28 | 1303.00588678 | 1338.85239212 | 655.23029298 |
| V29 | 668.48042280 | 655.23029298 | 387.41117794 |

Table 7.38

Linear Discriminant functions

| | <i>S. soubrense</i> | <i>S. damnosum s.s</i> | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
|----------|---------------------|------------------------|-----------------------|---------------------|
| CONSTANT | -280.89972742 | -260.34685396 | -314.02902408 | -292.11819108 |
| V4 | -0.01999528 | 0.03898369 | -0.02685458 | -0.04163823 |
| V6 | 0.42873824 | 0.40578646 | 0.38885047 | 0.37081618 |
| V9 | 0.98508531 | 0.69130814 | 1.08663561 | 0.77167116 |
| V18 | 0.15593835 | 0.18779172 | 0.17451851 | 0.21464193 |
| V20 | 0.11933462 | 0.14444330 | 0.14422052 | 0.14794828 |
| V23 | -0.09037097 | -0.11213230 | -0.12149516 | -0.02259459 |
| V27 | -0.02194982 | -0.11961641 | -0.09813606 | -0.05145069 |
| V28 | -0.34703442 | -0.22974707 | -0.31081350 | -0.39701843 |
| V29 | -0.27889690 | -0.38307706 | -0.16981360 | -0.15883762 |

7.3.3 DISCUSSION

The discriminant statistics presented in this section reveal that there is extensive multivariate morphometric differentiation between adult females of the four species of the *S. damnosum* complex for which samples were available from Togo and Benin.

The species which was most successfully discriminated relative to the other taxa, both using the species-pair statistics and in the overall analysis was *S. damnosum s.s.*. The maximum overlap of this species with any other was with *S. squamosum*, which was about 9% phenetic overlap. The influence of size variation on discrimination of *S. damnosum s.s.* was greater than for the other species, but even so, the size-free canonical roots were still larger when *S. damnosum s.s.* was involved than for most other species-pair discriminant analyses. The samples of *S. damnosum s.s.* used were smaller flies than the other species, but also they were a different shape, independent of size. Generally, the relative size of the antenna was the main morphological feature characterising *S. damnosum s.s.*, a finding in concordance with previous morphological studies of the *S. damnosum* complex (e.g. Garms 1978, Dang and Peterson 1980). The species also had consistently paler wing tufts than the other species, with the important exception of *S. squamosum*, to which it is closest morphologically.

Simulium squamosum is phenetically the next most isolated species, being closest to *S. damnosum s.s.*. The ability to discriminate between this species and the members of the *S. sanctipauli* subcomplex will be an important aid in the further understanding of the relative vectorial importance of the different species in Togo and Benin (see e.g., Cheke and Denke 1988). Previous morphological methods of

identification using thorax/antennal ratios found considerable overlap between these groups (Garms and Cheke 1985).

Simulium soubrense 'Beffa' and *S. sanctipauli* 'Djodji' show the greatest phenetic overlap, over 10%. The species-pair discriminant analysis also indicated that size differences were important in the between-species variation which had been found so that more extensive sampling is likely to reveal greater overlap than was found in this analysis. However, the discrimination between these species is greater than previous methods, which have not distinguished between the two, therefore it may be possible to collect more information as to the relative vectorial importance of these species. The restricted geographic range of *S. sanctipauli* 'Djodji' (Chapter two, Surtees *et al.* 1988), also means that the phenetic overlap between it and *S. soubrense* 'Beffa' may not be of critical importance, because they tend not to be found sympatrically.

Table 7.39 summarises the four methods of allocation used in the overall discriminant analysis: forced allocation with and without adjusting the prior probabilities of species membership using the fly's wing tuft colour, and typicality probability allocation with and without subsequent allocation of overlapping flies using the appropriate species-pair statistics. The relatively small sample sizes of the four species accounts for over 10% of flies overlapping using typicality probability allocation, because the approximate confidence intervals for each fly's distance to each of the species are broad. The subsequent species-pair allocation improves this. The largest number of correct allocations was using forced allocation with adjusted prior probabilities, although the effect of adjusting the prior probabilities is very slight. The smallest number of incorrect

allocations was obtained using typicality probability allocation. This method is more conservative than forced allocation, but considering the small sample sizes used to calculate the discriminant statistics this caution is well justified.

To conclude, the discriminant analyses presented in this section show that the four species of the *S. damnosum* complex which were examined from Togo and Benin can be successfully identified, although the rate of correct classification varies according to species. The ability to identify *S. damnosum* s.s., vector of the more debilitating form of onchocerciasis is clearly important, as is the ability to identify *S. squamosum*. The characters used in these analysis, and the use of multivariate statistical methods is without doubt an improvement over current morphological methods, and will be futher refined once larger samples have been obtained of each species, so that a wider range of variation, including seasonal size variation is sampled. The major limitation of the Togo and Benin statistics is that no *S. yahense* were available for analysis, despite its presence in Togo (Table 2.1). If account is to be taken of *S. yahense*, then the 'global' statistics developed in Chapter eight should be used, even though this assumes that *S. yahense* is the same morphologically in the east as in the west.

Table 7.39

Comparison of four methods of allocation for Togo and Benin

| | Forced ¹ | Forced ² | Typicality ³ | Typicality ⁴ |
|-------------|---------------------|---------------------|-------------------------|-------------------------|
| Correct | 174 | 176 | 154 | 166 |
| Incorrect | 16 | 14 | 10 | 10 |
| Overlapping | na | na | 20 | 8 |
| Atypical | na | na | 6 | 6 |

¹Forced allocation without adjusted priors

²Forced allocation with adjusted priors

³Typicality probability without subsequent species pair allocation of overlapping flies

⁴Typicality probability with subsequent species pair allocation of overlapping flies

7.4 RESULTS AND DISCUSSION, WESTERN AREA

7.4.1 SPECIES PAIR DISCRIMINATION

Samples from seven taxa were available from the area west of the Volta Lake, making the situation more complex than in Togo and Benin. For the purposes of the species pair statistics the single sample of *S. damnosum* s.s. was pooled with the five *S. sirbanum* samples into an artificial category called 'Savanna', because they were morphometrically very close (see section 7.4.1.1). Also OCP regards both species as dangerous vectors of onchocerciasis, to be controlled whenever either is found.

The single sample of *S. soubrense* 'B' was not pooled with the samples of *S. soubrense* even though it was for the overall analysis (see section 7.4.2). This was justified because of the chromosomal distinctiveness of this new species (Chapter three, Post 1986), even though morphometrically it was not distinctive (Chapter six).

7.4.1.1 Discrimination of Simulium damnosum and S. sirbanum

The squared Mahalanobis' distance between species using the 25 character set was 9.01 (unbiased $D^2=7.88$, $e_{act}=0.0802$) with 12 misallocated flies using resubstitution ($e_{res}=0.0583$). A stepwise discriminant analysis produced a nine character subset which gave a Mahalanobis' squared distance of 7.7 (unbiased $D^2=7.32$, $e_{act}=0.088$) with 18 flies misallocated using resubstitution ($e_{res}=0.087$).

The dimension reduction technique described in section 7.2.2 resulted in a seven character subset:

[V4, V9, V14, V16, V17, V23, V29]

i.e. one thorax, two antennal, two wing and two leg characters with a Mahalanobis' squared distance of 7.17 (unbiased $D^2=6.89$, $e_{act}=0.095$) and 17 misallocated flies using resubstitution ($e_{res}=0.083$), one *S.*

damnosum s.s. classified as *S. sirbanum* and 16 *S. sirbanum* classified as *S. damnosum s.s.*

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in 12 misallocated flies (5.83%), one atypical fly (0.49%), 15 overlapping flies (7.28%) and 178 correctly allocated flies (86.4%).

The null hypothesis of equal wing tuft colouration was not rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test. The prior probabilities of species membership for each wing tuft colouration category were therefore not calculated.

The null hypothesis of equal dispersion was not rejected at $p<0.001$ using the likelihood ratio test, legitimising the use of the pooled dispersion matrix.

The standardised canonical variate (Table 7.40) shows that the characters of most importance in discrimination are the positively loading characters: thorax width, tibia length 1, and basitarsus length 2, and the negatively loading characters: wing length 2 and antennal length 1.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

$$[0.41, 0.33, 0.17, 0.41, 0.41, 0.42, 0.42]$$

and accounted for 73.0% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 0.8498 to 0.5814 showing that size has considerable influence on the already relatively poor discrimination between these species.

The mean vectors are shown as Table 7.41, the pooled within-species dispersion matrix as Table 7.42 and the linear discriminant functions as Table 7.43.

To conclude, there is significant multivariate morphometric differentiation between *S. damnosum s.l.* and *S. sirbanum*, however this differentiation is relatively small, and is influenced to a considerable extent by size variation. Therefore, although the seven character subset classified correctly in over 86% of cases, this is probably an optimistic estimate, and can be expected to perform less well once a greater range of temporal and geographic variation within each species has been sampled. Operationally it is not important that these species cannot be differentiated very well as both are regarded as dangerous vectors of the more debilitating savanna form of onchocerciasis, and OCP controls both species whenever they are found.

Table 7.40

Standardised Canonical Variate for *S. damnosum s.s.* and *S. sirbanum*

| Character | Canonical Variate |
|--------------------|-------------------|
| Thorax Width | 0.8793 |
| Antennal Length 1 | -0.8197 |
| Antennal Segment 7 | 0.2192 |
| Wing Length 1 | -0.5568 |
| Wing Length 2 | -0.9491 |
| Tibia Length 1 | 0.8777 |
| Basitarsus Length | 1.0627 |

Table 7.41

Mean Vectors for species *S. damnosum s.s.* and *S. sirbanum*

| | <i>S. damnosum s.s.</i> | <i>S. sirbanum</i> |
|---------------------|-------------------------|--------------------|
| Thorax Width | 938.33485714 | 846.98794719 |
| Antennal Length 1 | 252.66428571 | 253.14943820 |
| Antennal Segment 7 | 36.89000000 | 35.70921348 |
| Wing Length 1 | 751.35428571 | 698.86112360 |
| Wing Length 2 | 457.56000000 | 432.46247191 |
| Tibia Length 1 | 716.56285714 | 657.28988764 |
| Basitarsus Length 2 | 330.73392857 | 301.03398876 |

Table 7.42

Pooled within-species dispersion matrix for species *S. damnosum s.s.* and *S. sirbanum*

| Character | V4 | V9 | V14 | V16 |
|-----------|---------------|---------------|--------------|---------------|
| V4 | 3334.50882078 | 532.59237158 | 45.66197091 | 2220.92410122 |
| V9 | 532.59237158 | 210.61198152 | 20.20367118 | 385.07727646 |
| V14 | 45.66197091 | 20.20367118 | 8.44149260 | 30.64210469 |
| V16 | 2220.92410122 | 385.07727646 | 30.64210469 | 2005.02023363 |
| V17 | 1481.87749590 | 272.14810905 | 21.90185856 | 1191.90521719 |
| V23 | 1894.67718138 | 335.59108748 | 30.47224600 | 1499.65516118 |
| V29 | 953.28978365 | 177.40024425 | 14.60045960 | 754.20342956 |
| Character | V17 | V23 | V29 | |
| V4 | 1481.87749590 | 1894.67718138 | 953.28978365 | |
| V9 | 272.14810905 | 335.59108748 | 177.40024425 | |
| V14 | 21.90185856 | 30.47224600 | 14.60045960 | |
| V16 | 1191.90521719 | 1499.65516118 | 754.20342956 | |
| V17 | 922.23770839 | 978.54598652 | 495.96160806 | |
| V23 | 978.54598652 | 1386.32722338 | 637.02149738 | |
| V29 | 495.96160806 | 637.02149738 | 357.89091733 | |

Table 7.43

Linear Discriminant functions for species *S. damnosum* s.s. and *S. sirbanum*

| | <i>S. damnosum</i> s.s. | <i>S. sirbanum</i> |
|----------|-------------------------|--------------------|
| CONSTANT | -234.18007423 | -211.54025531 |
| V4 | -0.08108991 | -0.11697217 |
| V9 | 0.52340679 | 0.67496271 |
| V14 | 1.91770284 | 1.71718024 |
| V16 | 0.03924445 | 0.07018309 |
| V17 | -0.27948955 | -0.19882914 |
| V23 | 0.57646803 | 0.52100165 |
| V29 | 0.08097309 | -0.05162960 |

7.4.1.2 Discrimination of 'Savanna' and *S. sanctipauli*

The squared Mahalanobis' distance between species using the 25 character set was 77.16 (unbiased $D^2=68.77$, $e_{act}<0.0001$) with no misallocated flies using resubstitution ($e_{res}=0.0$). A stepwise discriminant analysis produced a 12 character set, but the dimension reduction technique described in section 7.2.2 reduced this to a three character subset:

[V9,V17,V29]

i.e. one antennal, one wing and one leg character with a Mahalanobis' squared distance of 48.28 (unbiased $D^2=47.47$, $e_{act}=0.0003$) and no misallocated flies using resubstitution ($e_{res}=0.0$).

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in no misallocated flies, four atypical flies (1.66%), no overlapping flies and 237 correctly allocated flies (98.3%).

The null hypothesis of equal wing tuft colouration was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test. The prior probabilities of species membership for each wing tuft colouration category were calculated and are shown as Table 7.44. From this table it is very clear that this character alone would classify most flies

correctly, and when the prior probabilities were adjusted according to each fly's wing tuft colouration, the number of flies misallocated remained zero.

The null hypothesis of equal dispersion was not rejected at $p < 0.001$ using the likelihood ratio test, legitimising the use of the pooled dispersion matrix.

The standardised canonical variate (Table 7.45) shows that the characters of most importance in discrimination are the positively loading character antennal length 1, and the negatively loading character wing length 2, with *S. sanctipauli* lying at the positive end of this vector having relatively larger antennae.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

$$[0.51, 0.61, 0.60]$$

and accounted for 77.4% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 6.0438 to 3.2476 showing that size has some influence on the discrimination between these species, but that shape differences are important in the absence of size variation.

The mean vectors are shown as Table 7.46, the pooled within-species dispersion matrix as Table 7.47 and the discriminant functions as Table 7.48.

To conclude, there is significant multivariate morphometric differentiation between 'savanna' and *S. sanctipauli*. This differentiation is a combination of size and shape variation but in the absence of size variation, shape differences discriminate between this

species pair very well. Therefore, the three character set can be expected to classify correctly in over 98% of cases.

Adjusting the prior probabilities according to a fly's wing tuft colouration did not improve error rate, and so need only be used in cases of doubt following typicality probability allocation.

Table 7.44

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------|-----------------------|
| | 'Savanna' | <i>S. sanctipauli</i> |
| 1 | 1.0000 | 0.0000 |
| 2 | 1.0000 | 0.0000 |
| 3 | 0.0004 | 0.9996 |
| 4 | 0.0000 | 1.0000 |
| 5 | 0.0000 | 1.0000 |

Table 7.45

Standardised Canonical Variate for 'Savanna' and *S. sanctipauli*

| Character | Canonical Variate |
|-------------------|-------------------|
| Antennal Length 1 | 2.6164 |
| Wing Length 2 | -1.0412 |
| Basitarsus Length | 0.6315 |

Table 7.46

Mean Vectors for species *S. sanctipauli* and *S. squamosum*

| | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
|---------------------|-----------------------|---------------------|
| Antennal Length 1 | 253.08349515 | 336.06857143 |
| Wing Length 2 | 435.87378641 | 457.41942857 |
| Basitarsus Length 2 | 305.07087379 | 344.64857143 |

Table 7.47

Pooled within-species dispersion matrix for species *S. sanctipauli* and *S. squamosum*

| Character | V9 | V17 | V29 |
|-----------|--------------|--------------|--------------|
| V9 | 203.87903478 | 252.59119768 | 157.58476736 |
| V17 | 252.59119768 | 920.92884199 | 533.82209912 |
| V29 | 157.58476736 | 533.82209912 | 427.91009486 |

Table 7.48

Linear Discriminant functions for species 'Savanna' and *S. sanctipauli*

| | 'Savanna' | <i>S. sanctipauli</i> |
|----------|---------------|-----------------------|
| CONSTANT | -176.65358555 | -294.73381294 |
| V9 | 0.96290252 | 1.52092797 |
| V17 | 0.00537202 | -0.22632739 |
| V29 | 0.35162620 | 0.52766237 |

7.4.1.3 Discrimination of 'Savanna' and *S. soubrense* 'B' -

The squared Mahalanobis' distance between species using the 25 character set was 33.65 (unbiased $D^2=29.88$, $e_{act}=0.0031$) with no misallocated flies using resubstitution ($e_{res}=0.0$). Stepwise discriminant analysis applied to this character set produced an eight character subset with a D^2 of 32.15 (unbiased $D^2=30.7$, $e_{act}=0.0027$) and two misallocated flies using resubstitution ($e_{res}=0.0086$). The dimension reduction technique described in section 7.2.2 further reduced this to a five character subset:

[V3,V9,V13,V19,V29]

i.e. one thorax, two antennal, one wing and one leg character with a Mahalanobis' squared distance between species of 28.34 (unbiased $D^2=27.61$, $e_{act}=0.0043$) and two misallocated flies using resubstitution ($e_{res}=0.0086$).

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in one misallocated fly (0.43%), one atypical fly (0.43%), two overlapping flies (0.86%) and 230 correctly allocated flies (98.3%).

The null hypothesis of equal wing tuft colouration was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test. The prior probabilities of species membership for each wing tuft colouration category were calculated and are shown as Table 7.49. When the prior probabilities were adjusted according to each fly's wing tuft colouration, two flies were misallocated ($e_{res}=0.0086$).

The null hypothesis of equal dispersion was not rejected at $p<0.001$ using the likelihood ratio test, legitimising the use of the pooled dispersion matrix.

The standardised canonical variate (Table 7.50) shows that the contrast between the positively loading character antennal length 1, and the negatively loading character wing width 2, discriminates between these species, with *S. soubrense* 'B' lying at the positive end of this vector having relatively larger antennae.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

$$[0.47, 0.41, 0.35, 0.49, 0.50]$$

and accounted for 67.4% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 3.0111 to 3.0097 showing that size has virtually no influence on the discrimination between these species.

The mean vectors are shown as Table 7.51, the pooled within-species dispersion matrix as Table 7.52 and the linear discriminant functions as Table 7.53.

To conclude, there is significant multivariate morphometric differentiation between 'savanna' and *S. soubrense* 'B'. This differentiation is entirely shape variation involving the relative length of the antenna. Therefore, the three character set can be expected to classify correctly in over 98% of cases.

Adjusting the prior probabilities according to a fly's wing tuft colouration did not improve error rate, and so it is recommended that this should not be done except in cases of doubt.

Table 7.49

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------|-------------------------|
| | 'Savanna' | <i>S. soubrense</i> 'B' |
| 1 | 1.0000 | 0.0000 |
| 2 | 0.9450 | 0.0550 |
| 3 | 0.0003 | 0.9997 |
| 4 | 0.0000 | 1.0000 |
| 5 | 0.0000 | 1.0000 |

Table 7.50

Standardised Canonical Variate for 'Savanna' and *S. soubrense* 'B'

| Character | Canonical Variate |
|--------------------|-------------------|
| Thorax Length | -0.7091 |
| Antennal Length 1 | 1.0633 |
| Antennal Segment 6 | 0.7730 |
| Wing Width 2 | -0.9619 |
| Basitarsus Length | 0.8611 |

Table 7.51

Mean Vectors for species 'Savanna' and *S. soubrense* 'B'

| | 'Savanna' | <i>S. soubrense</i> 'B' |
|---------------------|---------------|-------------------------|
| Thorax Length | 618.73154272 | 586.83394286 |
| Antennal Length 1 | 253.08349515 | 299.88214286 |
| Antennal Segment 6 | 35.88174757 | 44.95000000 |
| Wing Width 2 | 1383.34075243 | 1349.70892857 |
| Basitarsus Length 2 | 305.07087379 | 315.69107143 |

Table 7.52

Pooled within-species dispersion matrix for species 'Savanna' and *S. soubrense* 'B'

| Character | V3 | V9 | V13 |
|-----------|---------------|---------------|--------------|
| V3 | 2083.38014492 | 337.07104861 | 57.99456940 |
| V9 | 337.07104861 | 216.20681015 | 22.08221096 |
| V13 | 57.99456940 | 22.08221096 | 8.86393005 |
| V19 | 2892.10382875 | 702.64485893 | 105.74630357 |
| V29 | 763.69444208 | 171.88295908 | 25.70746545 |
| Character | V19 | V29 | |
| V3 | 2892.10382875 | 763.69444208 | |
| V9 | 702.64485893 | 171.88295908 | |
| V13 | 105.74630357 | 25.70746545 | |
| V19 | 7074.88590766 | 1563.10241270 | |
| V29 | 1563.10241270 | 444.75467892 | |

Table 7.53

Linear Discriminant functions for species 'Savanna' and *S. soubrense* 'B'

| | 'Savanna' | <i>S. soubrense</i> 'B' |
|----------|---------------|-------------------------|
| CONSTANT | -185.24947818 | -234.34607400 |
| V3 | 0.02935689 | -0.05145444 |
| V9 | 0.72768806 | 0.99544177 |
| V13 | 0.93113997 | 1.91417183 |
| V19 | 0.13848980 | 0.07799207 |
| V29 | -0.18625321 | 0.02871013 |

7.4.1.4 Discrimination of 'Savanna' and *S. soubrense*

The squared Mahalanobis' distance between species using the 25

character set was 25.89 (unbiased $D^2=24.03$, $e_{act}=0.0071$) with six misallocated flies using resubstitution ($e_{res}=0.0165$). A stepwise discriminant analysis produced an eight character subset with a D^2 of 25.37 (unbiased $D^2=24.25$, $e_{act}=0.069$). The dimension reduction technique described in section 7.2.2 reduced this to a five character subset:

[V4,V10,V13,V20,V29]

i.e. one thorax, two antennal, one wing and one leg character with a Mahalanobis' squared distance of 17.82 (unbiased $D^2=17.53$, $e_{act}=0.0182$) between species and seven misallocated flies using resubstitution ($e_{res}=0.0192$), all seven of which were *S. soubrense* classified into *S. sirbanum*.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in six misallocated flies (1.65%), two atypical flies (0.55%), one overlapping fly (0.28%) and 355 correctly allocated flies (97.5%).

The null hypothesis of equal wing tuft colouration was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test. The prior probabilities of species membership for each wing tuft colouration category were calculated and are shown as Table 7.54. When the prior probabilities were adjusted according to each fly's wing tuft colouration, eight flies were misallocated ($e_{res}=0.022$).

The null hypothesis of equal dispersion was rejected at $p<0.001$ using the likelihood ratio test, but for the statistical and practical reasons given in section 7.2.5, the dispersion matrices were still pooled.

The standardised canonical variate (Table 7.55) shows that the most important contrast in discrimination is between antennal length

2. and thorax width. *S. soubrense* is at the positive side of this vector, having relatively larger antenna than 'savanna'.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.48,0.44,0.33,0.48,0.49]

and accounted for 74.6% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 4.4028 to 3.1025 showing that size has little influence on the discrimination between these species.

The mean vectors are shown as Table 7.56, the pooled within-species dispersion matrix as Table 7.57, and the linear discriminant functions as Table 7.58.

To conclude, there is significant multivariate morphometric differentiation between savanna and *S. soubrense*. This differentiation is mainly shape variation, so the five character set can be expected to classify correctly in over 97% of cases.

Adjusting the prior probabilities according to a fly's wing tuft colouration did not improve error rate, and so it is recommended that this should not be done except in cases of doubt.

Table 7.54

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------|---------------------|
| | 'Savanna' | <i>S. soubrense</i> |
| 1 | 0.8887 | 0.1113 |
| 2 | 0.4913 | 0.5087 |
| 3 | 0.1046 | 0.8954 |
| 4 | 0.0139 | 0.9861 |
| 5 | 0.0017 | 0.9983 |

Table 7.55

Standardised Canonical Variate for 'Savanna' and *S. soubrense*

| Character | Canonical Variate |
|--------------------|-------------------|
| Thorax Width | -1.1602 |
| Antennal Length 2 | 1.3912 |
| Antennal Segment 6 | 0.9875 |
| Wing Length 3 | 0.7299 |
| Basitarsus Length | -0.1533 |

Table 7.56

Mean Vectors for species 'Savanna' and *S. soubrense*

| | 'Savanna' | <i>S. soubrense</i> |
|---------------------|---------------|---------------------|
| Thorax Width | 859.40403204 | 876.48680506 |
| Antennal Length 2 | 380.12970874 | 456.40784810 |
| Antennal Segment 6 | 35.88174757 | 46.11544304 |
| Wing Length 3 | 1407.11145631 | 1504.37734177 |
| Basitarsus Length 2 | 305.07087379 | 328.69841772 |

Table 7.57

Pooled within-species dispersion matrix for species 'Savanna' and *S. soubrense*

| Character | V4 | V10 | V13 |
|-----------|---------------|---------------|--------------|
| V4 | 4116.58167146 | 1154.38579563 | 87.36060217 |
| V10 | 1154.38579563 | 703.06753937 | 53.53254960 |
| V13 | 87.36060217 | 53.53254960 | 10.43394721 |
| V20 | 5130.77823361 | 1597.06880493 | 123.04797006 |
| V29 | 1292.57184921 | 426.39566358 | 34.96686560 |
| Character | V20 | V29 | |
| V4 | 5130.77823361 | 1292.57184921 | |
| V10 | 1597.06880493 | 426.39566358 | |
| V13 | 123.04797006 | 34.96686560 | |
| V20 | 8131.90646643 | 1860.02049606 | |
| V29 | 1860.02049606 | 507.57891532 | |

Table 7.58

Linear Discriminant functions for species 'Savanna' and *S. soubrense*

| | 'Savanna' | <i>S. soubrense</i> |
|----------|---------------|---------------------|
| CONSTANT | -148.15174068 | -199.77811381 |
| V4 | -0.02663396 | -0.10241882 |
| V10 | 0.25651160 | 0.38364221 |
| V13 | 1.21771013 | 1.91060765 |
| V20 | 0.22571626 | 0.25587704 |
| V29 | -0.45765225 | -0.48316777 |

7.4.1.5 Discrimination of 'Savanna' and *S. squamosum*

The squared Mahalanobis' distance between species using the 25 character set was 18.14 (unbiased $D^2=16.52$, $e_{act}=0.0211$) with five misallocated flies using resubstitution ($e_{res}=0.0171$). Stepwise discriminant analysis produced an eight character subset with a D^2 between species of 17.17 (unbiased $D^2=16.16$, $e_{act}=0.022$) and ten misallocated flies ($e_{res}=0.034$). The dimension reduction technique described in section 7.2.2 resulted in a seven character subset:

[V4,V6,V9,V20,V27,V28,V29]

i.e. one thorax, one head, one antennal, one wing and three leg characters with a Mahalanobis' squared distance of 13.28 (unbiased $D^2=12.91$, $e_{act}=0.0362$) and ten misallocated flies using resubstitution ($e_{res}=0.034$), eight 'savanna' classified into *S. squamosum* and two *S. squamosum* classified into 'savanna'.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in five misallocated flies (1.71%), four atypical flies (1.37%), six overlapping fly (2.05%) and 278 correctly allocated flies (94.9%).

The null hypothesis of equal wing tuft colouration was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test. The prior probabilities of species membership for each wing tuft colouration category were calculated and are shown as Table 7.59. When the prior probabilities were adjusted according to each fly's wing tuft colouration, six flies were misallocated ($e_{res}=0.021$).

The null hypothesis of equal dispersion was rejected at $p<0.001$ using the likelihood ratio test, but for the statistical and practical reasons given in section 7.2.5, the dispersion matrices were still pooled.

The standardised canonical variate (Table 7.60) shows that the most important contrast in discrimination is between the positively loading characters: wing length 3, and basitarsus length 2 and the negatively loading characters: thorax width and tibia length 2.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.39,0.38,0.29,0.39,0.40,0.40,0.39]

and accounted for 85.8% of pooled within-species variation. When the

scores along this vector were introduced into the model as a covariable the canonical root fell from 2.7919 to 2.4606 showing that size has little influence on the discrimination between these species.

The mean vectors are shown as Table 7.61, the pooled within-species dispersion matrix as Table 7.62 and the linear discriminant functions as Table 7.63.

To conclude, there is significant multivariate morphometric differentiation between 'savanna' and *S. squamosum*. This differentiation is mainly shape variation, so the seven character set can be expected to classify correctly in nearly 95% of cases.

Adjusting the prior probabilities according to a fly's wing tuft colouration improved error rate, and so it is recommended that this should be done in cases of doubt following typicality probability allocation.

Table 7.59

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------|---------------------|
| | 'Savanna' | <i>S. squamosum</i> |
| 1 | 0.9013 | 0.0987 |
| 2 | 0.1873 | 0.8127 |
| 3 | 0.0058 | 0.9942 |
| 4 | 0.0001 | 0.9999 |
| 5 | 0.0000 | 1.0000 |

Table 7.60

Standardised Canonical Variate for 'Savanna' and *S. squamosum*

| Character | Canonical Variate |
|-------------------|-------------------|
| Thorax Width | -1.5111 |
| Head Width | -0.2167 |
| Antennal Length 1 | 0.5834 |
| Wing Length 3 | 1.4802 |
| Femur Length 2 | 0.9878 |
| Tibia Length 2 | -2.0050 |
| Basitarsus Length | 1.4755 |

Table 7.61

Mean Vectors for species 'Savanna' and *S. squamosum*

| | 'Savanna' | <i>S. squamosum</i> |
|---------------------|---------------|---------------------|
| Thorax Width | 859.40403204 | 855.21798621 |
| Head Width | 775.71054175 | 795.31917241 |
| Antennal Length 1 | 253.08349515 | 276.18275862 |
| Wing Length 3 | 1407.11145631 | 1506.67091954 |
| Femur Length 2 | 631.47961165 | 658.14896552 |
| Tibia Length 2 | 597.82776699 | 612.96413793 |
| Basitarsus Length 2 | 305.07087379 | 329.43793103 |

Table 7.62

Pooled within-species dispersion matrix for species 'Savanna' and *S. squamosum*

| Character | V4 | V6 | V9 | V20 |
|-----------|---------------|---------------|---------------|---------------|
| V4 | 4081.94577108 | 2731.89999834 | 589.35997597 | 4759.88112753 |
| V6 | 2731.89999834 | 2324.38954083 | 451.37162330 | 3534.85810881 |
| V9 | 589.35997597 | 451.37162330 | 228.47447774 | 799.74397450 |
| V20 | 4759.88112753 | 3534.85810881 | 799.74397450 | 7084.34530099 |
| V27 | 2431.88941630 | 1775.76543610 | 392.03597295 | 3224.99834846 |
| V28 | 2324.16176789 | 1714.74046939 | 378.45810520 | 3113.27041767 |
| V29 | 1210.31479711 | 901.31965086 | 201.69990676 | 1637.40522063 |
| Character | V27 | V28 | V29 | |
| V4 | 2431.88941630 | 2324.16176789 | 1210.31479711 | |
| V6 | 1775.76543610 | 1714.74046939 | 901.31965086 | |
| V9 | 392.03597295 | 378.45810520 | 201.69990676 | |
| V20 | 3224.99834846 | 3113.27041767 | 1637.40522063 | |
| V27 | 1708.94166521 | 1611.01703297 | 827.81038517 | |
| V28 | 1611.01703297 | 1604.65839616 | 796.96262868 | |
| V29 | 827.81038517 | 796.96262868 | 455.92922705 | |

Table 7.63

Linear Discriminant functions for species 'Savanna' and *S. squamosum*

| | 'Savanna' | <i>S. squamosum</i> |
|----------|---------------|---------------------|
| CONSTANT | -196.45847171 | -238.41977587 |
| V4 | -0.20494826 | -0.29125591 |
| V6 | 0.24599641 | 0.22986491 |
| V9 | 0.70930704 | 0.82471657 |
| V20 | 0.21088635 | 0.26732438 |
| V27 | 0.23528375 | 0.31893918 |
| V28 | -0.17417372 | -0.35421631 |
| V29 | -0.46702989 | -0.24350292 |

7.4.1.6 Discrimination of 'Savanna' and *S. yahense*

The squared Mahalanobis' distance between species using the 25 character set was 46.59 (unbiased $D^2=42.55$, $e_{act}=0.0006$) with one misallocated fly using resubstitution ($e_{res}=0.0033$). A stepwise discriminant analysis produced a 12 character subset with a D^2 of 45.21 (unbiased $D^2=43.25$, $e_{act}=0.0005$) and one misallocated fly

($e_{res}=0.0033$). The dimension reduction technique described in section 7.2.2 resulted in a four character subset:

[V4,V9,V20,V29]

i.e. one thorax, one antennal, one wing and one leg character with a Mahalanobis' squared distance of 30.1 (unbiased $D^2=29.6$, $e_{act}=0.0033$) and one misallocated flies using resubstitution ($e_{res}=0.0033$), which was a 'savanna' fly classified into *S. yahense*.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in no misallocated flies, four atypical flies (1.32%), no overlapping flies and 298 correctly allocated flies (98.68%).

The null hypotheses of equal wing tuft colouration and equal abdominal setal colouration were both rejected at $p<0.001$ using Wilcoxon two-sample rank sum tests. The prior probabilities of species membership for each wing tuft colouration category were calculated and are shown as Table 7.64. When the prior probabilities were adjusted according to each fly's wing tuft colouration, no flies were misallocated ($e_{res}=0.0$). However, when the prior probabilities were adjusted for each abdominal setal colouration category, four flies were misallocated ($e_{res}=0.0132$), the four being *S. yahense* flies with abdominal setal colouration category one (Chapter four).

The null hypothesis of equal dispersion was rejected at $p<0.001$ using the likelihood ratio test, but for the statistical and practical reasons given in section 7.2.5, the dispersion matrices were still pooled.

The standardised canonical variate (Table 7.65) shows that the most important characters in discrimination are antennal length 1 and the negatively loading character thorax width.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.52,0.43,0.52,0.52]

and accounted for 83.1% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 6.5698 to 3.4387 showing that size has some influence on the discrimination between these species, but that shape differences are important in its absence.

The mean vectors are shown as Table 7.66, the pooled within-species dispersion matrix as Table 7.67 and the linear discriminant functions as Table 7.68.

To conclude, there is significant multivariate morphometric differentiation between 'savanna' and *S. yahense*. This differentiation is a combination of size and shape variation, with shape variation being more important, so the four character set can be expected to classify correctly in nearly 99% of cases.

Adjusting the prior probabilities according to a fly's wing tuft colouration improved error rate slightly, but since the fly which was allocated correctly as a result was actually atypical, it is recommended that this should be done only in cases of doubt (i.e. only on 'typical' overlapping flies). Adjusting prior probabilities for each abdominal setal colouration category did not improve error rate, so it is not recommended that this should be used.

Table 7.64

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------|-------------------|
| | 'Savanna' | <i>S. yahense</i> |
| 1 | 1.0000 | 0.0000 |
| 2 | 1.0000 | 0.0000 |
| 3 | 0.0712 | 0.9288 |
| 4 | 0.0000 | 1.0000 |
| 5 | 0.0000 | 1.0000 |

Table 7.65

Standardised Canonical Variate for 'Savanna' and *S. yahense*

| Character | Canonical Variate |
|-------------------|-------------------|
| Thorax Width | -1.5496 |
| Antennal Length 1 | 2.2583 |
| Wing Length 3 | 0.5025 |
| Basitarsus Length | 0.9848 |

Table 7.66

Mean Vectors for species 'Savanna' and *S. yahense*

| | 'Savanna' | <i>S. yahense</i> |
|---------------------|---------------|-------------------|
| Thorax Width | 859.40403204 | 906.62228125 |
| Antennal Length 1 | 253.08349515 | 319.98281250 |
| Wing Length 3 | 1407.11145631 | 1576.95807292 |
| Basitarsus Length 2 | 305.07087379 | 352.44218750 |

Table 7.67

Pooled within-species dispersion matrix for species 'Savanna' and *S. yahense*

| Character | V4 | V9 | V20 | V29 |
|-----------|---------------|--------------|---------------|---------------|
| V4 | 4344.69214529 | 669.65652051 | 5096.98270101 | 1285.79911384 |
| V9 | 669.65652051 | 251.12671008 | 890.65435174 | 225.75435193 |
| V20 | 5096.98270101 | 890.65435174 | 7566.57036711 | 1739.46501848 |
| V29 | 1285.79911384 | 225.75435193 | 1739.46501848 | 485.64772294 |

Table 7.68

Linear Discriminant functions for species 'Savanna' and *S. yahense*

| | 'Savanna' | <i>S. yahense</i> |
|----------|---------------|-------------------|
| CONSTANT | -168.33314123 | -253.58470268 |
| V4 | -0.12202774 | -0.24453803 |
| V9 | 0.68764496 | 1.04177862 |
| V20 | 0.23796683 | 0.26141911 |
| V29 | -0.22073570 | -0.04745514 |

7.4.1.7 Discrimination of Simulium sanctipauli and S. soubrense 'B'

The squared Mahalanobis' distance between species using the 25 character set was 25.49 (unbiased $D^2=14.63$, $e_{act}=0.0279$) with no misallocated flies using resubstitution ($e_{res}=0.0$). A stepwise discriminant analysis produced an eleven character subset with a D^2 of 22.68 (unbiased $D^2=18.22$, $e_{act}=0.0164$) and no misallocated flies ($e_{res}=0.0$). The dimension reduction technique described in section 7.2.2 resulted in a six character subset:

[V4,V10,V15,V16,V20,V24]

i.e. one thorax, two antennal, two wing and one leg character with a Mahalanobis' squared distance of 12.76 (unbiased $D^2=11.3$, $e_{act}=0.0464$) and no misallocated flies using resubstitution ($e_{res}=0.0$).

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in no misallocated flies, one atypical fly (1.59%), eight overlapping fly (12.7%) and 54 correctly allocated flies (85.7%).

The null hypothesis of equal wing tuft colouration was not rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test, so the prior probabilities of species membership for each wing tuft colouration category were not calculated.

The null hypothesis of equal dispersion was not rejected at $p<0.001$ using the likelihood ratio test legitimising the use of the pooled dispersion matrix.

The standardised canonical variate (Table 7.69) shows that the most important character in discrimination is the positively loading character: basitarsus length 1.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.44,0.37,0.21,0.45,0.45,0.46]

and accounted for 66.2% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 3.2546 to 1.1461 indicating that size is important in discriminating between these two species.

The mean vectors are shown as Table 7.70, the pooled within-species dispersion matrix as Table 7.71 and the linear discriminant functions as Table 7.72.

To conclude, there is significant multivariate morphometric differentiation between *S. sanctipauli* and *S. soubrense* 'B', however much of this differentiation is size variation, although there is still significant discrimination in its absence. The six character subset

classified nearly 86% of flies correctly, but once a wider range of temporal and geographic variation has been sampled, this figure may be optimistic.

Table 7.69

Standardised Canonical Variate for *S. sanctipauli* and *S. soubrense* 'B'

| Character | Canonical Variate |
|--------------------|-------------------|
| Thorax Width | -0.7089 |
| Antennal Length 2 | 0.8290 |
| Antennal Segment 8 | 0.8452 |
| Wing Length 1 | 0.4008 |
| Wing Length 3 | -0.6764 |
| Basitarsus Length | 1.2457 |

Table 7.70

Mean Vectors for species *S. sanctipauli* and *S. soubrense* 'B'

| | <i>S. sanctipauli</i> | <i>S. soubrense</i> 'B' |
|---------------------|-----------------------|-------------------------|
| Thorax Width | 906.40608000 | 854.03575714 |
| Antennal Length 2 | 501.27771429 | 441.74571429 |
| Antennal Segment 8 | 50.20228571 | 42.91285714 |
| Wing Length 1 | 771.73714286 | 707.60142857 |
| Wing Length 3 | 1512.93442857 | 1422.18357143 |
| Basitarsus Length 1 | 459.10628571 | 416.44285714 |

Table 7.71

Pooled within-species dispersion matrix for species *S. sanctipauli* and *S. soubrense* 'B'

| Character | V4 | V10 | V15 |
|-----------|---------------|---------------|---------------|
| V4 | 2521.06007214 | 673.54413359 | 39.50850533 |
| V10 | 673.54413359 | 499.30943775 | 34.56310370 |
| V15 | 39.50850533 | 34.56310370 | 8.80771129 |
| V16 | 1511.12786324 | 437.44422857 | 22.64939859 |
| V20 | 2730.06872634 | 879.84030464 | 63.61045508 |
| V24 | 801.37675347 | 266.71257780 | 11.87441424 |
| Character | V16 | V20 | V24 |
| V4 | 1511.12786324 | 2730.06872634 | 801.37675347 |
| V10 | 437.44422857 | 879.84030464 | 266.71257780 |
| V15 | 22.64939859 | 63.61045508 | 11.87441424 |
| V16 | 1512.06383044 | 2378.01545902 | 708.64661171 |
| V20 | 2378.01545902 | 5435.19611691 | 1248.19125639 |
| V24 | 708.64661171 | 1248.19125639 | 430.94194735 |

Table 7.72

Linear Discriminant functions for species *S. sanctipauli* and *S. soubrense* 'B'

| | <i>S. sanctipauli</i> | <i>S. soubrense</i> 'B' |
|----------|-----------------------|-------------------------|
| CONSTANT | -351.35623515 | -277.83430877 |
| V4 | -0.10026713 | -0.05527611 |
| V10 | 0.44030726 | 0.36060156 |
| V15 | 3.06561061 | 2.42182736 |
| V16 | 0.03974933 | 0.01122305 |
| V20 | 0.04088571 | 0.06894794 |
| V24 | 0.71104427 | 0.56107719 |

7.4.1.8 Discrimination of *Simulium sanctipauli* and *S. soubrense*

The squared Mahalanobis' distance between species using the 25 character set was 11.57 (unbiased $D^2=10.0$, $e_{act}=0.0569$) with nine misallocated flies using resubstitution ($e_{res}=0.0415$). A stepwise discriminant analysis produced an eleven character subset with a D^2 of 11.05 (unbiased $D^2 =10.36$, $e_{act}=0.0538$) and 12 misallocated flies

($e_{res}=0.062$). The dimension reduction technique described in section 7.2.2 resulted in a five character subset:

[V11,V15,V19,V20,V27]

i.e. two antennal, two wing and one leg character with a Mahalanobis' squared distance of 8.2 (unbiased $D^2=7.94$, $e_{act}=0.0794$) and 12 misallocated flies using resubstitution ($e_{res}=0.06$), of which 10 were *S. soubrense* misclassified as *S. sanctipauli* and two were *S. sanctipauli* misclassified as *S. soubrense*.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in eight misallocated flies (4.15%), two atypical flies (1.04%), ten overlapping flies (5.18%) and 173 correctly allocated flies (89.6%).

The null hypothesis of equal wing tuft colouration was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test, so the prior probabilities of species membership for each wing tuft colouration category were calculated and are shown as Table 7.73. When the prior probabilities of species membership were adjusted according to a fly's wing tuft colouration, the number of flies misallocated rose to 18 ($e_{res}=0.093$).

The null hypothesis of equal dispersion was not rejected at $p<0.001$ using the likelihood ratio test, legitimising the use of the pooled dispersion matrix.

The standardised canonical variate (Table 7.74) shows that the most important character in discrimination is the negatively loading character wing length 3, with femur length 2 having some opposite influence

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.34,0.38,0.48,0.50,0.50]

and accounted for 68.8% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 1.23 to 1.1461 indicating that size is not important in discriminating between these species.

The mean vectors are shown as Table 7.75, the pooled within-species dispersion matrix as Table 7.77 and the linear discriminant functions as Table 7.78.

To conclude, there is significant multivariate morphometric differentiation between *S. sanctipauli* and *S. soubrense*, although this differentiation is not great. The five character subset can be expected to classify nearly 90% of flies correctly when it is known *a priori* that just this species pair can be expected.

Adjusting prior probabilities according to a fly's wing tuft colouration did not improve error rate, so it is not recommended that this should be done.

Table 7.73

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------------------|---------------------|
| | <i>S. sanctipauli</i> | <i>S. soubrense</i> |
| 1 | 0.1022 | 0.8978 |
| 2 | 0.2105 | 0.7895 |
| 3 | 0.3843 | 0.6157 |
| 4 | 0.5937 | 0.4063 |
| 5 | 0.7738 | 0.2262 |

Table 7.74

Standardised Canonical Variate for *S. sanctipauli* and *S. soubrense*

| Character | Canonical Variate |
|--------------------|-------------------|
| Antennal Segment 4 | 0.5997 |
| Antennal Segment 8 | 0.5237 |
| Wing Width 2 | 0.7372 |
| Wing Length 3 | -2.1817 |
| Femur Length 2 | 1.2937 |

Table 7.75

Mean Vectors for species *S. sanctipauli* and *S. soubrense*

| | <i>S. sanctipauli</i> | <i>S. soubrense</i> |
|--------------------|-----------------------|---------------------|
| Antennal Segment 4 | 49.91885714 | 42.41113924 |
| Antennal Segment 8 | 50.20228571 | 44.09848101 |
| Wing Width 2 | 1488.56614286 | 1423.85350633 |
| Wing Length 3 | 1512.93442857 | 1504.37734177 |
| Femur Length 2 | 692.31428571 | 663.60835443 |

Table 7.76

Pooled within-species dispersion matrix for species *S. sanctipauli* and *S. soubrense*

| Character | V11 | V15 | V19 |
|-----------|---------------|---------------|---------------|
| V11 | 17.42241439 | 8.62668673 | 136.96191809 |
| V15 | 8.62668673 | 12.15275525 | 141.42825524 |
| V19 | 136.96191809 | 141.42825524 | 7812.75567753 |
| V20 | 165.28271757 | 162.66671418 | 7280.48278354 |
| V27 | 71.62593229 | 72.04564640 | 3060.16200252 |
| Character | V20 | V27 | |
| V11 | 165.28271757 | 71.62593229 | |
| V15 | 162.66671418 | 72.04564640 | |
| V19 | 7280.48278354 | 3060.16200252 | |
| V20 | 8833.97372629 | 3535.85005757 | |
| V27 | 3535.85005757 | 1612.55914885 | |

Table 7.77

Linear Discriminant functions for species *S. sanctipauli* and *S. soubrense*

| | <i>S. sanctipauli</i> | <i>S. soubrense</i> |
|----------|-----------------------|---------------------|
| CONSTANT | -181.19268736 | -155.23678949 |
| V11 | 0.73449596 | 0.39603997 |
| V15 | 1.69442424 | 1.33750211 |
| V19 | 0.10332047 | 0.08028381 |
| V20 | -0.07219432 | -0.00559610 |
| V27 | 0.28322704 | 0.19409267 |

7.4.1.9 Discrimination of *Simulium sanctipauli* and *S. squamosum*

The squared Mahalanobis' distance between species using the 25 character set was 80.49 (unbiased $D^2=63.4$, $e_{act}<0.0001$) with no misallocated flies using resubstitution ($e_{res}=0.0$). A stepwise discriminant analysis produced a nine character subset which gave a D^2 of 72.48 (unbiased $D^2=66.44$, $e_{act}<0.0001$) and no misallocated flies ($e_{res}=0.0$). The dimension reduction technique described in section 7.2.2 resulted in a two character subset:

[V10,V20]

i.e. one antennal and one wing character with a Mahalanobis' squared distance of 41.42 (unbiased $D^2=40.73$, $e_{act}=0.0007$) and no misallocated flies using resubstitution ($e_{res}=0.0$).

Allocation using typicality probability²² of species membership with atypicality defined at $\alpha=0.01$ resulted in no misallocated flies, one atypical fly (0.82%), no overlapping flies and 121 correctly allocated flies (99.2%).

The null hypothesis of equal wing tuft colouration was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test, so the prior probabilities of species membership for each wing tuft colouration category were calculated and are shown as Table 7.78. When the prior

probabilities of species membership were adjusted according to a fly's wing tuft colouration, no flies were misallocated.

The null hypothesis of equal dispersion was not rejected at $p < 0.001$ using the likelihood ratio test legitimising the use of the pooled dispersion matrix.

The standardised canonical variate (Table 7.79) shows that most discrimination is due to antennal length 2 with an opposite effect from wing length 3.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.71, 0.71]

and accounted for 88.9% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 8.62 to 6.9401 indicating that size has some influence in discrimination, but that shape variation is much more important.

The mean vectors are shown as Table 7.80, the pooled within-species dispersion matrix as Table 7.81 and the linear discriminant functions as Table 7.82.

To conclude, there is significant multivariate morphometric differentiation between *S. sanctipauli* and *S. squamosum*, and the two character subset derived in this analysis can be expected to correctly classify over 99% of flies when it is known *a priori* that just this species pair can be expected.

Adjusting prior probabilities according to a fly's wing tuft colouration should be done in cases of doubt following typicality probability allocation.

Table 7.78

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------------------|---------------------|
| | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
| 1 | 0.0008 | 0.9992 |
| 2 | 0.0207 | 0.9793 |
| 3 | 0.3623 | 0.6377 |
| 4 | 0.9386 | 0.0614 |
| 5 | 0.9976 | 0.0024 |

Table 7.79

Standardised Canonical Variate for *S. sanctipauli* and *S. squamosum*

| Character | Canonical Variate |
|-------------------|-------------------|
| Antennal Length 2 | 3.3654 |
| Wing Length 3 | -1.2291 |

Table 7.80

Mean Vectors for species *S. sanctipauli* and *S. squamosum*

| | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
|-------------------|-----------------------|---------------------|
| Antennal Length 2 | 501.27771429 | 405.58896552 |
| Wing Length 3 | 1512.93442857 | 1506.67091954 |

Table 7.81

Pooled within-species dispersion matrix for species *S. sanctipauli* and *S. squamosum*

| Character | V10 | V20 |
|-----------|---------------|---------------|
| V10 | 547.65873687 | 1494.64465864 |
| V20 | 1494.64465864 | 6709.02769555 |

Table 7.82

Linear Discriminant functions for species *S. sanctipauli* and *S. squamosum*

| | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
|----------|-----------------------|---------------------|
| CONSTANT | -233.40249585 | -180.56918886 |
| V10 | 0.76497296 | 0.32574574 |
| V20 | 0.05508573 | 0.15200366 |

7.4.1.10 Discrimination of *Simulium sanctipauli* and *S. yahense*

The squared Mahalanobis' distance between species using the 25 character set was 14.58 (unbiased $D^2=11.64$ $e_{act}=0.044$) with four misallocated flies using resubstitution ($e_{res}=0.0305$). A stepwise discriminant analysis produced an eleven character subset with a D^2 of 12.91 (unbiased $D^2=11.71$, $e_{act}=0.0435$) and five misallocated flies ($e_{res}=0.0382$). The dimension reduction technique described in section 7.2.2 resulted in a six character subset:

[V3,V9,V11,V17,V20,V22]

i.e. one thorax, two antennal, two wing and one leg character with a Mahalanobis' squared distance of 10.07 (unbiased $D^2=9.52$, $e_{act}=0.0615$) and six misallocated flies using resubstitution ($e_{res}=0.046$), all six of which were *S. yahense* misclassified as *S. sanctipauli*.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in five misallocated flies (3.82%), four atypical flies (3.05%), six overlapping flies (4.58%) and 116 correctly allocated flies (88.55%).

The null hypotheses of equal wing tuft colouration and equal abdominal setal colouration were both rejected at $p<0.001$ using Wilcoxon two-sample rank sum tests. The prior probabilities of species membership for each wing tuft colouration category were calcu-

lated and are shown as Table 7.83. When the prior probabilities of species membership were adjusted according to a fly's wing tuft colouration, the number of flies misallocated was seven ($e_{res} = 0.053$). When the prior probabilities of species membership were adjusted according to a fly's abdominal setal colouration category, the number of flies misallocated was four (the four *S. yahense* with abdominal setal colouration category one).

The null hypothesis of equal dispersion was not rejected at $p < 0.001$ using the likelihood ratio test, legitimising the use of the pooled dispersion matrix.

The standardised canonical variate (Table 7.84) shows that the most important contrast of characters is between femur length 1 and wing length 3.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

$$[0.40, 0.39, 0.37, 0.43, 0.43, 0.44]$$

and accounted for 77.3% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 1.7747 to 1.7412 indicating that size is not important in discriminating between these species.

The mean vectors are shown as Table 7.85, the pooled within-species dispersion matrix as Table 7.86 and the linear discriminant functions as Table 7.87.

To conclude, there is significant multivariate morphometric differentiation between *S. sanctipauli* and *S. yahense*, mainly involving shape variation. The six character subset can be expected to classify nearly 89% of flies correctly when it is known *a priori* that just this species pair can be expected.

Adjusting prior probabilities according to a fly's wing tuft colouration did not improve greatly improve error rate, so it is recommended that this should be done only in cases of doubt following typicality probability allocation, likewise with abdominal setal colouration.

Table 7.83

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------------------|-------------------|
| | <i>S. sanctipauli</i> | <i>S. yahense</i> |
| 1 | 0.9991 | 0.0009 |
| 2 | 0.9934 | 0.0066 |
| 3 | 0.9524 | 0.0476 |
| 4 | 0.7265 | 0.2735 |
| 5 | 0.2609 | 0.7391 |

Table 7.84

Standardised Canonical Variate for *S. sanctipauli* and *S. yahense*

| Character | Canonical Variate |
|--------------------|-------------------|
| Thorax Length | 0.4583 |
| Antennal Length 1 | 0.6552 |
| Antennal Segment 4 | 0.7283 |
| Wing Length 2 | -0.8580 |
| Wing Length 3 | -1.8563 |
| Femur Length 1 | 1.1017 |

Table 7.85

Mean Vectors for species *S. sanctipauli* and *S. yahense*

| | <i>S. sanctipauli</i> | <i>S. yahense</i> |
|--------------------|-----------------------|-------------------|
| Thorax Length | 657.81216000 | 650.18355000 |
| Antennal Length 1 | 336.06857143 | 319.98281250 |
| Antennal Segment 4 | 49.91885714 | 45.00166667 |
| Wing Length 2 | 457.41942857 | 471.03875000 |
| Wing Length 3 | 1512.93442857 | 1576.95807292 |
| Femur Length 1 | 657.87428571 | 657.64000000 |

Table 7.86

Pooled within-species dispersion matrix for species *S. sanctipauli* and *S. yahense*

| Character | V3 | V9 | V11 |
|-----------|---------------|---------------|---------------|
| V3 | 2101.72893060 | 476.38310941 | 130.43457686 |
| V9 | 476.38310941 | 295.53507418 | 56.21923638 |
| V11 | 130.43457686 | 56.21923638 | 22.27158828 |
| V17 | 1090.34712382 | 392.28948301 | 93.66614091 |
| V20 | 3119.57653853 | 1071.46097048 | 279.49498206 |
| V22 | 1493.70925395 | 559.31001329 | 130.80799823 |
| Character | V17 | V20 | V22 |
| V3 | 1090.34712382 | 3119.57653853 | 1493.70925395 |
| V9 | 392.28948301 | 1071.46097048 | 559.31001329 |
| V11 | 93.66614091 | 279.49498206 | 130.80799823 |
| V17 | 994.05007937 | 2279.75222448 | 1170.04748438 |
| V20 | 2279.75222448 | 7856.66628688 | 3479.19583477 |
| V22 | 1170.04748438 | 3479.19583477 | 1884.81091827 |

Table 7.87

Linear Discriminant functions for species *S. sanctipauli* and *S. yahense*

| | <i>S. sanctipauli</i> | <i>S. yahense</i> |
|----------|-----------------------|-------------------|
| CONSTANT | -239.03455218 | -244.46837753 |
| V3 | 0.09601691 | 0.06611952 |
| V9 | 1.31761714 | 1.21212558 |
| V11 | -2.13424984 | -2.55398160 |
| V17 | -0.09305999 | -0.01291173 |
| V20 | 0.21241801 | 0.27220438 |
| V22 | -0.30426702 | -0.38037825 |

7.4.1.11 Discrimination of *Simulium soubrense* 'B' and *S. soubrense*

The squared Mahalanobis' distance between species using the 25 character set was 5.67 (unbiased $D^2=4.87$, $e_{act}=0.135$) with 19 misallocated flies using resubstitution ($e_{res}=0.102$). A stepwise discriminant analysis produced a ten character subset with a D^2 of 4.56 (unbiased $D^2=4.29$, $e_{act}=0.1509$) and 27 misallocated flies ($e_{res}=0.145$). The dimension reduction technique described in section 7.2.2 resulted in an eight character subset:

[V3,V4,V16,V17,V20,V21,V22,V28]

i.e. two thorax, four wing and two leg character with a Mahalanobis' squared distance of 4.09 (unbiased $D^2=3.89$, $e_{act}=0.162$) and 26 misallocated flies using resubstitution ($e_{res}=0.1398$), of which two were *S. soubrense* 'B' misclassified as *S. soubrense*, and 24 *S. soubrense* misclassified as *S. soubrense* 'B'.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in 17 misallocated flies (9.14%), seven atypical flies (3.76%), 29 overlapping flies (15.6%) and 133 correctly allocated flies (71.51%).

The null hypothesis of equal wing tuft colouration was not rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test, therefore

the prior probabilities of species membership for each wing tuft colouration category were not calculated.

The null hypothesis of equal dispersion was not rejected at $p < 0.001$ using the likelihood ratio test, legitimising the use of the pooled dispersion matrix.

The standardised canonical variate (Table 7.88) shows that the most important contrast of characters is between wing length 3, femur length 1, wing length 2, and the negatively loading characters tibia length 2, :thorax width and wing length 1.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.33,0.37,0.37,0.37,0.37,0.25,0.37,0.38]

and accounted for 82.2% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 0.5287 to 0.4281 indicating that size is important in discriminating between these species. Such a small canonical root suggests that the between species variation is in reality very small, and perhaps biologically unmeaningful (Campbell 1982).

The mean vectors are shown as Table 7.89, the pooled within-species dispersion matrix as Table 7.90 and the linear discriminant functions as Table 7.91.

To conclude, there is significant multivariate morphometric differentiation between *S. soubrense* 'B' and *S. soubrense*, but this is very slight in comparison to other species-pair discriminant analyses. It is not recommended that this species pair be separated using this seven character subset.

Table 7.88

Standardised Canonical Variate for *S. soubrense* and *S. soubrense* 'B'

| Character | Canonical Variate |
|--------------------|-------------------|
| Thorax Length | 0.6167 |
| Thorax Width | -1.5308 |
| Wing Length 1 | -1.2355 |
| Wing Length 2 | 1.2174 |
| Wing Length 3 | 1.9824 |
| Radial Hair Number | 0.4037 |
| Femur Length 1 | 1.5214 |
| Tibia Length 2 | -2.2162 |

Table 7.89

Mean Vectors for species *S. soubrense* and *S. soubrense* 'B'

| | <i>S. soubrense</i> | <i>S. soubrense</i> 'B' |
|--------------------|---------------------|-------------------------|
| Thorax Length | 586.83394286 | 628.90365570 |
| Thorax Width | 854.03575714 | 876.48680506 |
| Wing Length 1 | 707.60142857 | 738.09341772 |
| Wing Length 2 | 419.78142857 | 448.12481013 |
| Wing Length 3 | 1422.18357143 | 1504.37734177 |
| Radial Hair Number | 13.07142857 | 14.82278481 |
| Femur Length 1 | 607.09285714 | 640.53417722 |
| Tibia Length 2 | 600.24000000 | 622.53569620 |

Table 7.90

Pooled within-species dispersion matrix for species *S. soubrense* and *S. soubrense* 'B'

| Characte | V3 | V4 | V16 | V17 |
|----------|---------------|---------------|---------------|---------------|
| V3 | 2753.67086144 | 2540.46154488 | 2109.08913798 | 1346.92465004 |
| V4 | 2540.46154488 | 3623.86125944 | 2712.06032687 | 1715.20579338 |
| V16 | 2109.08913798 | 2712.06032687 | 2802.81829727 | 1600.95871383 |
| V17 | 1346.92465004 | 1715.20579338 | 1600.95871383 | 1101.65924939 |
| V20 | 3902.89663612 | 5142.61995803 | 4687.19056899 | 2830.88356890 |
| V21 | 64.50138627 | 101.31871863 | 88.23615673 | 53.73658601 |
| V22 | 1705.81167474 | 2203.91076242 | 1906.70059038 | 1187.42862995 |
| V28 | 1655.36329864 | 2128.43777672 | 1889.07338437 | 1179.67898734 |
| Characte | V20 | V21 | V22 | V28 |
| V3 | 3902.89663612 | 64.50138627 | 1705.81167474 | 1655.36329864 |
| V4 | 5142.61995803 | 101.31871863 | 2203.91076242 | 2128.43777672 |
| V16 | 4687.19056899 | 88.23615673 | 1906.70059038 | 1889.07338437 |
| V17 | 2830.88356890 | 53.73658601 | 1187.42862995 | 1179.67898734 |
| V20 | 9271.31758194 | 150.65254580 | 3552.95675473 | 3530.32897496 |
| V21 | 150.65254580 | 8.70051694 | 69.42897417 | 68.36891029 |
| V22 | 3552.95675473 | 69.42897417 | 1674.31753269 | 1521.45783500 |
| V28 | 3530.32897496 | 68.36891029 | 1521.45783500 | 1538.07225801 |

Table 7.91

Linear Discriminant functions for species *S. soubrense* and *S. soubrense* 'B'

| | <i>S. soubrense</i> | <i>S. soubrense</i> 'B' |
|----------|---------------------|-------------------------|
| CONSTANT | -145.13327813 | -150.85930038 |
| V3 | -0.10416705 | -0.08126909 |
| V4 | 0.09896614 | 0.04785658 |
| V16 | -0.00356864 | -0.04991006 |
| V17 | -0.21647810 | -0.14537616 |
| V20 | 0.02961330 | 0.06952539 |
| V21 | -2.63733323 | -2.36590739 |
| V22 | 0.10765730 | 0.17999523 |
| V28 | 0.47859779 | 0.36632780 |

7.4.1.12 Discrimination of *Simulium soubrense* 'B' and *S. squamosum*

The squared Mahalanobis' distance between species using the 25 character set was 34.67 (unbiased $D^2=26.69$ $e_{act}=0.0049$) with no mis-allocated flies using resubstitution ($e_{res}=0.0$). A stepwise

discriminant analysis produced a 12 character subset with a D^2 of 31.57 (unbiased $D^2 = 27.94$, $e_{act} = 0.0041$) and no misallocated flies ($e_{res} = 0.0$). The dimension reduction technique described in section 7.2.2 resulted in an four character subset:

[V4,V10,V17,V20]

i.e. one thorax, one antennal, and two wing characters with a Mahalanobis' squared distance of 20.75 (unbiased $D^2 = 19.83$, $e_{act} = 0.01299$) and one misallocated fly using resubstitution ($e_{res} = 0.0087$), which was a *S. soubrense* 'B' misclassified as *S. squamosum*.

Allocation using typicality probability of species membership with atypicality defined at $\alpha = 0.01$ resulted in one misallocated fly (0.87%), one atypical fly (0.87%), one overlapping flies (0.87%) and 112 correctly allocated flies (97.39%).

The null hypothesis of equal wing tuft colouration was rejected at $p < 0.001$ using a Wilcoxon two-sample rank sum test, therefore the prior probabilities of species membership for each wing tuft colouration category were calculated and are shown as Table 7.92. Adjusting the prior probability of species membership according to a fly's wing tuft colouration resulted in a single misallocated fly ($e_{res} = 0.0087$).

The null hypothesis of equal dispersion was not rejected at $p < 0.001$ using the likelihood ratio test so the pooled dispersion matrix was used.

The standardised canonical variate (Table 7.93) shows that the most important contrast of characters is between the positively loading character antennal length 2 and the negatively loading characters wing length 2, and wing length 3.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.51,0.48,0.51,0.51]

and accounted for 86.2% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 3.8905 to 3.7107 indicating that size is of little importance in discriminating between these species.

The mean vectors are shown as Table 7.94, the pooled within-species dispersion matrix as Table 7.95 and the linear discriminant functions as Table 7.96.

To conclude, there is significant multivariate morphometric differentiation between *S. soubrense* 'B' and *S. squamosum*. The four character subset can be expected to allocate correctly in over 97% of cases when it is known *a priori* that just this species pair can be expected. The lack of influence of size variation in discrimination implies that once a wider range of temporal and geographic variation is sampled, the error rate should not be substantially worse.

Adjusting the prior probabilities of species membership according to the fly's wing tuft colouration did not improve error rate, so it is not recommended that this should be done.

Table 7.92

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-------------------------|---------------------|
| | <i>S. soubrense</i> 'B' | <i>S. squamosum</i> |
| 1 | 0.0570 | 0.9430 |
| 2 | 0.2275 | 0.7725 |
| 3 | 0.5894 | 0.4106 |
| 4 | 0.8750 | 0.1250 |
| 5 | 0.9715 | 0.0285 |

Table 7.93

Standardised Canonical Variate for *S. soubrense* 'B' and *S. squamosum*

| Character | Canonical Variate |
|-------------------|-------------------|
| Thorax Width | 0.6841 |
| Antennal Length 2 | 1.8070 |
| Wing Length 2 | -1.2721 |
| Wing Length 3 | -1.0838 |

Table 7.94

Mean Vectors for species *S. soubrense* 'B' and *S. squamosum*

| | <i>S. soubrense</i> 'B' | <i>S. squamosum</i> |
|-------------------|-------------------------|---------------------|
| Thorax Width | 854.03575714 | 855.21798621 |
| Antennal Length 2 | 441.74571429 | 405.58896552 |
| Wing Length 2 | 419.78142857 | 450.71724138 |
| Wing Length 3 | 1422.18357143 | 1506.67091954 |

Table 7.95

Pooled within-species dispersion matrix for species *S. soubrense* 'B' and *S. squamosum*

| Character | V4 | V10 | V17 | V20 |
|-----------|---------------|---------------|---------------|---------------|
| V4 | 3225.08076155 | 1016.36998911 | 1400.14442037 | 4194.91822849 |
| V10 | 1016.36998911 | 539.42068401 | 531.85084180 | 1540.33015320 |
| V17 | 1400.14442037 | 531.85084180 | 842.04421310 | 2093.11773874 |
| V20 | 4194.91822849 | 1540.33015320 | 2093.11773874 | 7289.52899841 |

Table 7.96

Linear Discriminant functions for species *S. soubrense* 'B' and *S. squamosum*

| | <i>S. soubrense</i> 'B' | <i>S. squamosum</i> |
|----------|-------------------------|---------------------|
| CONSTANT | -191.16952958 | -176.45339324 |
| V4 | -0.02693330 | -0.08204850 |
| V10 | 0.76679613 | 0.47157782 |
| V17 | -0.21570523 | -0.03357885 |
| V20 | 0.11050698 | 0.16390038 |

7.4.1.13 Discrimination of *Simulium soubrense* 'B' and *S. yahense*

The squared Mahalanobis' distance between species using the 25 character set was 13.05 (unbiased $D^2=10.27$ $e_{act}=0.0055$) with six misallocated flies using resubstitution ($e_{res}=0.0484$). A stepwise discriminant analysis produced an 11 character subset with a D^2 of 11.65 (unbiased $D^2=10.5$, $e_{act}=0.0053$) and six misallocated flies ($e_{res}=0.0484$). The dimension reduction technique described in section 7.2.2 resulted in an eight character subset:

[V4, V10, V17, V19, V20, V27, V28, V29]

i.e. one thorax, one antennal, three wing and three leg characters with a Mahalanobis' squared distance of 8.11 (unbiased $D^2=7.51$, $e_{act}=0.0853$) and eight misallocated flies using resubstitution ($e_{res}=0.065$), two *S. soubrense* 'B' into *S. yahense*, and six *S. yahense* into *S. soubrense* 'B'.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in six misallocated flies (4.84%), five atypical flies (4.03%), 10 overlapping flies (8.07%) and 103 correctly allocated flies (83.06%).

The null hypotheses of equal wing tuft colouration and equal abdominal setal colouration were both rejected at $p<0.001$ using Wilcoxon two-sample rank sum tests. The prior probabilities of species membership for each wing tuft colouration category were calcu-

lated and are shown as Table 7.97. Adjusting the prior probability of species membership according to a fly's wing tuft colouration resulted in two misallocated flies ($e_{res}=0.0161$). Adjusting the prior probabilities according to a fly's abdominal setal colouration resulted in four misallocated flies ($e_{res}=0.0323$), these being the four *S. yahense* with abdominal setal colouration category one (Chapter four).

The null hypothesis of equal dispersion was not rejected at $p<0.001$ using the likelihood ratio test so the pooled dispersion matrix was used, but the multivariate statistical test for skewness (Mardia 1970) showed some departure from normality, which could account for the relatively large proportion of atypical flies.

The standardised canonical variate (Table 7.98) shows that there are only two negatively loading characters, thorax width and tibia length 2, indicating that this discriminant vector may have a considerable size component.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.36,0.32,0.35,0.35,0.36,0.37,0.37,0.36]

and accounted for 87.4% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 1.4407 to 0.755 confirming the fact that size is of importance in discriminating between these species.

The mean vectors are shown as Table 7.99, the pooled within-species dispersion matrix as Table 7.100 and the linear discriminant functions as Table 7.101.

To conclude, there is significant multivariate morphometric differentiation between *S. soubrense* 'B' and *S. yahense*, but size variation contributes significantly to this discrimination. Therefore the eight character subset can be expected to allocate correctly about 83% of the time, but once a wider range of size variation has been sampled, it is likely that this estimate is optimistic.

Adjusting the prior probabilities of species membership according to the fly's wing tuft colouration improved error rate, so it is recommended that this should be done. Adjusting the prior probabilities according to abdominal setal colouration is subject to the extreme influence that this character imposes (Chapter nine) but the better performance of this character over the subset derived in this analysis means that its use is recommended.

Table 7.97

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-------------------------|-------------------|
| | <i>S. soubrense</i> 'B' | <i>S. yahense</i> |
| 1 | 1.0000 | 0.0000 |
| 2 | 0.9988 | 0.0012 |
| 3 | 0.9711 | 0.0289 |
| 4 | 0.5677 | 0.4323 |
| 5 | 0.0489 | 0.9511 |

Table 7.98

Standardised Canonical Variate for *S. soubrense* 'B' and *S. yahense*

| Character | Canonical Variate |
|-------------------|-------------------|
| Thorax Width | -1.6063 |
| Antennal Length 2 | 0.2239 |
| Wing Length 2 | 0.8375 |
| Wing Width 2 | 0.9903 |
| Wing Length 3 | 1.1733 |
| Femur Length 2 | 0.2859 |
| Tibia Length 2 | -1.7129 |
| Basitarsus Length | 0.9400 |

Table 7.99

Mean Vectors for species *S. soubrense* 'B' and *S. yahense*

| | <i>S. soubrense</i> 'B' | <i>S. yahense</i> |
|---------------------|-------------------------|-------------------|
| Thorax Width | 854.03575714 | 906.62228125 |
| Antennal Length 2 | 441.74571429 | 475.44625000 |
| Wing Length 2 | 419.78142857 | 471.03875000 |
| Wing Width 2 | 1349.70892857 | 1495.92739583 |
| Wing Length 3 | 1422.18357143 | 1576.95807292 |
| Femur Length 2 | 633.80142857 | 690.33750000 |
| Tibia Length 2 | 600.24000000 | 648.41500000 |
| Basitarsus Length 2 | 315.69107143 | 352.44218750 |

Table 7.100

Pooled within-species dispersion matrix for species *S. soubrense* 'B' and *S. yahense*

| Characte | V4 | V10 | V17 | V19 |
|----------|---------------|---------------|---------------|---------------|
| V4 | 3934.38975621 | 1599.54679697 | 1726.76572239 | 4858.91600003 |
| V10 | 1599.54679697 | 987.09609292 | 740.83766329 | 2055.55129618 |
| V17 | 1726.76572239 | 740.83766329 | 1052.89838519 | 2414.28330107 |
| V19 | 4858.91600003 | 2055.55129618 | 2414.28330107 | 7656.33734523 |
| V20 | 5065.53411486 | 2156.62348348 | 2468.48739038 | 7126.09981005 |
| V27 | 2589.69969903 | 1134.30120386 | 1296.31614133 | 3440.14839687 |
| V28 | 2376.82672350 | 1019.78553934 | 1208.88977213 | 3229.93999139 |
| V29 | 1183.15732295 | 511.09617308 | 614.36624934 | 1661.39299409 |
| Characte | V20 | V27 | V28 | V29 |
| V4 | 5065.53411486 | 2589.69969903 | 2376.82672350 | 1183.15732295 |
| V10 | 2156.62348348 | 1134.30120386 | 1019.78553934 | 511.09617308 |
| V17 | 2468.48739038 | 1296.31614133 | 1208.88977213 | 614.36624934 |
| V19 | 7126.09981005 | 3440.14839687 | 3229.93999139 | 1661.39299409 |
| V20 | 8460.19183903 | 3700.08764012 | 3485.06270144 | 1808.42069014 |
| V27 | 3700.08764012 | 2023.16538478 | 1825.05343279 | 935.82996297 |
| V28 | 3485.06270144 | 1825.05343279 | 1758.70078033 | 861.32035451 |
| V29 | 1808.42069014 | 935.82996297 | 861.32035451 | 493.23935991 |

Table 7.101

Linear Discriminant functions for species *S. soubrense* 'B' and *S. yahense*

| | <i>S. soubrense</i> 'B' | <i>S. yahense</i> |
|----------|-------------------------|-------------------|
| CONSTANT | -138.09186312 | -173.61086040 |
| V4 | -0.12195821 | -0.19099451 |
| V10 | 0.23775813 | 0.25632349 |
| V17 | -0.09095115 | -0.02952965 |
| V19 | 0.11991277 | 0.14636684 |
| V20 | 0.10357705 | 0.13332578 |
| V27 | -0.09082624 | -0.07476517 |
| V28 | 0.05108723 | -0.05401815 |
| V29 | 0.09895643 | 0.19821470 |

7.4.1.14 Discrimination of *Simulium soubrense* and *S. squamosum*

The squared Mahalanobis' distance between species using the 25 character set was 13.35 (unbiased $D^2=11.92$ $e_{act}=0.0421$) with nine misallocated flies using resubstitution ($e_{res}=0.0367$). A stepwise

discriminant analysis produced a ten character subset with a D^2 of 12.49 (unbiased $D^2 = 11.93$ $e_{act} = 0.0421$) and eight misallocated flies ($e_{res} = 0.0327$). The dimension reduction technique described in section 7.2.2 resulted in an eight character subset:

[V3,V4,V6,V10,V13,V24,V28,V29]

i.e. two thorax, one head, two antennal, and three leg characters with a Mahalanobis' squared distance of 11.92 (unbiased $D^2 = 11.48$, $e_{act} = 0.04152$) and ten misallocated flies using resubstitution ($e_{res} = 0.0408$), nine *S. soubrense* into *S. squamosum*, and one *S. squamosum* into *S. soubrense*.

Allocation using typicality probability of species membership with atypicality defined at $\alpha = 0.01$ resulted in four misallocated flies (1.63%), three atypical flies (1.22%), six overlapping flies (2.45%) and 232 correctly allocated flies (94.69%).

The null hypothesis of equal wing tuft colouration was rejected at $p < 0.001$ using a Wilcoxon two-sample rank sum test. The prior probabilities of species membership for each wing tuft colouration category were calculated and are shown as Table 7.102. Adjusting the prior probability of species membership according to a fly's wing tuft colouration resulted in a 11 misallocated flies ($e_{res} = 0.0449$).

The null hypothesis of equal dispersion was not rejected at $p < 0.001$ using the likelihood ratio test, so the pooled dispersion matrix was used.

The standardised canonical variate (Table 7.103) shows that antennal length 2 is important in discrimination, as well as thorax shape and leg length.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.34,0.37,0.37,0.34,0.25,0.38,0.38,0.38]

and accounted for 78.6% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 2.7528 to 2.6439 showing that size variation is of little importance relative to shape variation in discriminating between these species.

The mean vectors are shown as Table 7.105, the pooled within-species dispersion matrix as Table 7.106 and the linear discriminant functions as Table 7.107.

To conclude, there is significant multivariate morphometric differentiation between *S. soubrense* and *S. squamosum*, which is mainly shape variation. The eight character subset can be expected to allocate correctly in nearly 95% of the cases.

Adjusting the prior probabilities of species membership according to the fly's wing tuft colouration made the error rate worse, so it is not recommended that this should be done.

Table 7.102

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|---------------------|---------------------|
| | <i>S. soubrense</i> | <i>S. squamosum</i> |
| 1 | 0.3118 | 0.6882 |
| 2 | 0.4346 | 0.5654 |
| 3 | 0.5660 | 0.4340 |
| 4 | 0.6888 | 0.3112 |
| 5 | 0.7897 | 0.2103 |

Table 7.103

Standardised Canonical Variate for *S. soubrense* and *S. squamosum*

| Character | Canonical Variate |
|--------------------|-------------------|
| Thorax Length | -0.4038 |
| Thorax Width | 0.6438 |
| Head Width | -0.4703 |
| Antennal Length 2 | 1.6163 |
| Antennal Segment 6 | 0.6236 |
| Basitarsus Length | -0.9714 |
| Tibia Length 2 | 0.8248 |
| Basitarsus Length | -0.8872 |

Table 7.104

Mean Vectors for species *S. soubrense* and *S. squamosum*

| | <i>S. soubrense</i> | <i>S. squamosum</i> |
|---------------------|---------------------|---------------------|
| Thorax Length | 628.90365570 | 631.97175172 |
| Thorax Width | 876.48680506 | 855.21798621 |
| Head Width | 809.42156962 | 795.31917241 |
| Antennal Length 2 | 456.40784810 | 405.58896552 |
| Antennal Segment 6 | 46.11544304 | 39.08137931 |
| Basitarsus Length 1 | 435.45113924 | 436.35310345 |
| Tibia Length 2 | 622.53569620 | 612.96413793 |
| Basitarsus Length 2 | 328.69841772 | 329.43793103 |

Table 7.105

Pooled within-species dispersion matrix for species *S. soubrense* and *S. squamosum*

| Character | V3 | V4 | V6 | V10 |
|-----------|---------------|---------------|---------------|---------------|
| V3 | 2907.60441486 | 2664.11298272 | 1845.18036992 | 994.99138316 |
| V4 | 2664.11298272 | 3760.51158467 | 2433.10407715 | 1251.91824591 |
| V6 | 1845.18036992 | 2433.10407715 | 2052.78305406 | 997.75622128 |
| V10 | 994.99138316 | 1251.91824591 | 997.75622128 | 796.09547521 |
| V13 | 79.61103673 | 101.85343706 | 83.68869484 | 67.46957932 |
| V24 | 1243.10401386 | 1623.72048579 | 1162.11681398 | 660.69797682 |
| V28 | 1736.66642653 | 2191.73641037 | 1549.97195216 | 824.25232802 |
| V29 | 946.07944730 | 1234.12716620 | 891.09574800 | 498.13886574 |
| Character | V13 | V24 | V28 | V29 |
| V3 | 79.61103673 | 1243.10401386 | 1736.66642653 | 946.07944730 |
| V4 | 101.85343706 | 1623.72048579 | 2191.73641037 | 1234.12716620 |
| V6 | 83.68869484 | 1162.11681398 | 1549.97195216 | 891.09574800 |
| V10 | 67.46957932 | 660.69797682 | 824.25232802 | 498.13886574 |
| V13 | 12.79452738 | 57.26154752 | 69.73669632 | 43.90896423 |
| V24 | 57.26154752 | 906.30697719 | 1100.17549489 | 649.40660717 |
| V28 | 69.73669632 | 1100.17549489 | 1583.32221722 | 844.21730152 |
| V29 | 43.90896423 | 649.40660717 | 844.21730152 | 525.58342359 |

Table 7.106

Linear Discriminant functions for species *S. soubrense* and *S. squamosum*

| | <i>S. soubrense</i> | <i>S. squamosum</i> |
|----------|---------------------|---------------------|
| CONSTANT | -190.76465438 | -170.00160277 |
| V3 | -0.09902805 | -0.07312541 |
| V4 | -0.09027153 | -0.12610120 |
| V6 | 0.41966966 | 0.45519018 |
| V10 | 0.20337566 | 0.05350385 |
| V13 | 1.11867749 | 0.68023840 |
| V24 | -0.14500433 | -0.03337458 |
| V28 | 0.39341413 | 0.32216414 |
| V29 | -0.43487297 | -0.30099220 |

7.4.1.15 Discrimination of *Simulium soubrense* and *S. yahense*

The squared Mahalanobis' distance between species using the 25 character set was 5.03 (unbiased $D^2=4.51$ $e_{act}=0.1442$) with 34 mislocated flies using resubstitution ($e_{res}=0.134$). A stepwise

discriminant analysis produced a nine character subset with a D^2 of 4.28 (unbiased $D^2 = 4.12$, $e_{act} = 0.155$) and 35 misallocated flies ($e_{res} = 0.1378$). The dimension reduction technique described in section 7.2.2 resulted in an eight character subset:

[V6,V18,V19,V22,V24,V25,V26,V29]

i.e. one head, two wing and five leg characters with a Mahalanobis' squared distance of 4.2 (unbiased $D^2 = 4.05$, $e_{act} = 0.1572$) and 33 misallocated flies using resubstitution ($e_{res} = 0.1299$), 22 *S. soubrense* into *S. yahense*, and 11 *S. yahense* into *S. soubrense*.

Allocation using typicality probability of species membership with atypicality defined at $\alpha = 0.01$ resulted in 23 misallocated flies (9.06%), 10 atypical flies (3.94%), 24 overlapping flies (9.45%) and 197 correctly allocated flies (77.56%).

The null hypotheses of equal wing tuft colouration and equal abdominal setal colouration were both rejected at $p < 0.001$ using Wilcoxon two-sample rank sum tests. The prior probabilities of species membership for each wing tuft colouration category were calculated and are shown as Table 7.108. Adjusting the prior probability of species membership according to a fly's wing tuft colouration resulted in 34 misallocated flies ($e_{res} = 0.1339$). Adjusting the prior probabilities according to a fly's abdominal setal colouration resulted in four misallocated flies ($e_{res} = 0.0157$), these being the four *S. yahense* with abdominal setal colouration category one (Chapter four).

The null hypothesis of equal dispersion was not rejected at $p < 0.001$ using the likelihood ratio test legitimising the use of the pooled dispersion matrix

The standardised canonical variate (Table 7.108) shows that the main characters discriminating between this species pair are femur length 1, basitarsus length 1, and basitarsus length 2.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.35,0.35,0.35,0.36,0.37,0.36,0.34,0.36]

and accounted for 86.3% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 0.9954 to 0.807 showing that size is of some importance in discriminating between these species.

The mean vectors are shown as Table 7.109, the pooled within-species dispersion matrix as Table 7.110 and the linear discriminant functions as Table 7.111

To conclude, the eight character subset offers very poor discrimination between *S. soubrense* and *S. yahense*, and it is not recommended that they should be identified in this way.

Using the abdominal setal colour character gives a more satisfactory error rate, so it is recommended that this character be used on its own.

Table 7.107

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|---------------------|-------------------|
| | <i>S. soubrense</i> | <i>S. yahense</i> |
| 1 | 0.9864 | 0.0136 |
| 2 | 0.9405 | 0.0595 |
| 3 | 0.7747 | 0.2253 |
| 4 | 0.4279 | 0.5721 |
| 5 | 0.1399 | 0.8601 |

Table 7.108

Standardised Canonical Variate for *S. soubrense* and *S. yahense*

| Character | Canonical Variate |
|-------------------|-------------------|
| Head Width | -0.5451 |
| Wing Width 1 | -0.6656 |
| Wing Width 2 | 0.7507 |
| Femur Length 1 | -1.3216 |
| Basitarsus Length | 0.6647 |
| Tarsus Segment 2 | -1.3443 |
| Tarsus Segment 3 | 0.6487 |
| Basitarsus Length | 2.4347 |

Table 7.109

Mean Vectors for species *S. soubrense* and *S. yahense*

| | <i>S. soubrense</i> | <i>S. yahense</i> |
|---------------------|---------------------|-------------------|
| Head Width | 809.42156962 | 828.68378125 |
| Wing Width 1 | 1009.80537975 | 1039.91921875 |
| Wing Width 2 | 1423.85350633 | 1495.92739583 |
| Femur Length 1 | 640.53417722 | 657.64000000 |
| Basitarsus Length 1 | 435.45113924 | 462.27500000 |
| Tarsus Segment 2 | 167.56481013 | 172.72166667 |
| Tarsus Segment 3 | 125.46759494 | 131.40125000 |
| Basitarsus Length 2 | 328.69841772 | 352.44218750 |

Table 7.110

Pooled within-species dispersion matrix for species *S. soubrense* and *S. yahense*

| Character | V6 | V18 | V19 | V22 |
|-----------|---------------|---------------|---------------|---------------|
| V6 | 2124.90433379 | 2443.20135162 | 3350.45289437 | 1770.44352484 |
| V18 | 2443.20135162 | 4506.90502776 | 5215.71907686 | 2601.06808710 |
| V19 | 3350.45289437 | 5215.71907686 | 8374.65043572 | 3474.59832884 |
| V22 | 1770.44352484 | 2601.06808710 | 3474.59832884 | 1967.02331525 |
| V24 | 1224.49148122 | 1771.16381253 | 2419.45455484 | 1242.25677083 |
| V25 | 403.85449661 | 580.51226840 | 795.27150171 | 410.31138581 |
| V26 | 318.43274511 | 492.37409063 | 633.02918931 | 319.55345630 |
| V29 | 937.62187253 | 1344.89810519 | 1807.00430882 | 954.17537914 |
| Character | V24 | V25 | V26 | V29 |
| V6 | 1224.49148122 | 403.85449661 | 318.43274511 | 937.62187253 |
| V18 | 1771.16381253 | 580.51226840 | 492.37409063 | 1344.89810519 |
| V19 | 2419.45455484 | 795.27150171 | 633.02918931 | 1807.00430882 |
| V22 | 1242.25677083 | 410.31138581 | 319.55345630 | 954.17537914 |
| V24 | 956.85661109 | 296.57958863 | 228.08255886 | 687.79443268 |
| V25 | 296.57958863 | 116.88226380 | 85.98369376 | 228.53422402 |
| V26 | 228.08255886 | 85.98369376 | 81.78923625 | 174.71857078 |
| V29 | 687.79443268 | 228.53422402 | 174.71857078 | 558.47493550 |

Table 7.111

Linear Discriminant functions for species *S. soubrense* and *S. yahense*

| | <i>S. soubrense</i> | <i>S. yahense</i> |
|----------|---------------------|-------------------|
| CONSTANT | -186.57383097 | -189.29742336 |
| V6 | 0.44733902 | 0.42354367 |
| V18 | 0.11935613 | 0.09946490 |
| V19 | 0.07931896 | 0.09504821 |
| V22 | -0.17732607 | -0.23746916 |
| V24 | -0.28476846 | -0.24411637 |
| V25 | 1.35795592 | 1.10923261 |
| V26 | -0.54475594 | -0.40445037 |
| V29 | -0.43813258 | -0.24808245 |

7.4.1.16 Discrimination of *Simulium squamosum* and *S. yahense*

The squared Mahalanobis' distance between species using the 25 character set was 20.64 (unbiased $D^2=17.67$ $e_{act}=0.0178$) with no mis-allocated flies using resubstitution ($e_{res}=0.0$). A stepwise discriminant analysis produced an 11 character subset with a D^2 of

19.73 (unbiased $D^2 = 18.42$, $e_{act} = 0.016$) and no misallocated flies ($e_{res} = 0.0$). The dimension reduction technique described in section 7.2.2 resulted in an seven character subset:

[V9,V11,V15,V16,V22,V24,V29]

i.e. one head, two wing and five leg characters with a Mahalanobis' squared distance of 16.03 (unbiased $D^2 = 15.32$, $e_{act} = 0.0252$) and one misallocated fly using resubstitution ($e_{res} = 0.0055$), which was a *S. squamosum* misidentified as *S. yahense*.

Allocation using typicality probability of species membership with atypicality defined at $\alpha = 0.01$ resulted in one misallocated fly (0.55%), four atypical flies (2.19%), two overlapping flies (1.09%) and 176 correctly allocated flies (96.17%).

The null hypotheses of equal wing tuft colouration and equal abdominal setal colouration were both rejected at $p < 0.001$ using Wilcoxon two-sample rank sum tests. The prior probabilities of species membership for each wing tuft colouration category were calculated and are shown as Table 7.112. Adjusting the prior probability of species membership according to a fly's wing tuft colouration resulted in no misallocated flies ($e_{res} = 0.0$). Adjusting the prior probabilities according to a fly's abdominal setal colouration resulted in four misallocated flies ($e_{res} = 0.0323$), these being the four *S. yahense* with abdominal setal colouration category one (Chapter four).

The null hypothesis of equal dispersion was not rejected at $p < 0.001$ using the likelihood ratio test and so the pooled dispersion matrix was used.

The standardised canonical variate (Table 7.113) shows that antennal length 1 and basitarsus length 2 contrast with wing length 1 and the other leg measurements.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.38,0.32,0.34,0.40,0.40,0.40,0.40]

and accounted for 79.6% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 4.0427 to 3.0228 implying that size has some influence in discrimination, but that shape differences are more important.

The mean vectors are shown as Table 7.114, the pooled within-species dispersion matrix as Table 7.115 and the linear discriminant functions as Table 7.116.

To conclude, there is significant multivariate morphometric differentiation between *S. squamosum* and *S. yahense*, mainly involving shape differences. The seven character subset can be expected to allocate correctly over 96% of the time.

Adjusting the prior probabilities of species membership according to the fly's wing tuft colouration improved error rate, so it is recommended that this should be done. However, adjusting the prior probabilities according to abdominal setal colouration is subject to the extreme influence that this character imposes (Chapter nine) so it is recommended that it should be used only in cases of doubt, otherwise all *S. yahense* with pale abdominal setal colouration will be incorrectly allocated.

Table 7.112

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|---------------------|-------------------|
| | <i>S. squamosum</i> | <i>S. yahense</i> |
| 1 | 1.0000 | 0.0000 |
| 2 | 0.9998 | 0.0002 |
| 3 | 0.9238 | 0.0762 |
| 4 | 0.0341 | 0.9659 |
| 5 | 0.0001 | 0.9999 |

Table 7.113

Standardised Canonical Variate for *S. squamosum* and *S. yahense*

| Character | Canonical Variate |
|--------------------|-------------------|
| Antennal Length 1 | 1.7646 |
| Antennal Segment 4 | 0.7048 |
| Antennal Segment 8 | 0.4806 |
| Wing Length 1 | -0.8171 |
| Femur Length 1 | -0.8805 |
| Basitarsus Length | -0.6487 |
| Basitarsus Length | 1.0703 |

Table 7.114

Mean Vectors for species *S. squamosum* and *S. yahense*

| | <i>S. squamosum</i> | <i>S. yahense</i> |
|---------------------|---------------------|-------------------|
| Antennal Length 1 | 276.18275862 | 319.98281250 |
| Antennal Segment 4 | 36.13103448 | 45.00166667 |
| Antennal Segment 8 | 37.96965517 | 46.06083333 |
| Wing Length 1 | 734.04137931 | 770.69750000 |
| Femur Length 1 | 631.34344828 | 657.64000000 |
| Basitarsus Length 1 | 436.35310345 | 462.27500000 |
| Basitarsus Length 2 | 329.43793103 | 352.44218750 |

Table 7.115

Pooled within-species dispersion matrix for species *S. squamosum* and *S. yahense*

| Characte | V9 | V11 | V15 | V16 |
|----------|---------------|---------------|---------------|---------------|
| V9 | 308.74590209 | 46.21971990 | 45.24130861 | 792.98599417 |
| V11 | 46.21971990 | 16.27798641 | 8.43954142 | 147.95524793 |
| V15 | 45.24130861 | 8.43954142 | 10.92835261 | 129.31021791 |
| V16 | 792.98599417 | 147.95524793 | 129.31021791 | 3270.62386317 |
| V22 | 634.17394018 | 115.96479055 | 100.58862046 | 2402.34416125 |
| V24 | 408.80214036 | 76.17205426 | 61.18805632 | 1474.16430291 |
| V29 | 302.42489782 | 55.11946318 | 48.38336886 | 1158.70728493 |
| Characte | V22 | V24 | V29 | |
| V9 | 634.17394018 | 408.80214036 | 302.42489782 | |
| V11 | 115.96479055 | 76.17205426 | 55.11946318 | |
| V15 | 100.58862046 | 61.18805632 | 48.38336886 | |
| V16 | 2402.34416125 | 1474.16430291 | 1158.70728493 | |
| V22 | 2072.96679870 | 1195.38910756 | 927.55830177 | |
| V24 | 1195.38910756 | 836.71949537 | 600.89218237 | |
| V29 | 927.55830177 | 600.89218237 | 495.40067109 | |

Table 7.116

Linear Discriminant functions for species *S. squamosum* and *S. yahense*

| | <i>S. squamosum</i> | <i>S. yahense</i> |
|----------|---------------------|-------------------|
| CONSTANT | -145.00160013 | -186.98928625 |
| V9 | 0.75681927 | 1.00850247 |
| V11 | -0.80731730 | -0.33643093 |
| V15 | -0.22155901 | 0.14686223 |
| V16 | -0.17568891 | -0.23029386 |
| V22 | 0.00051998 | -0.07405566 |
| V24 | 0.23273406 | 0.15061693 |
| V29 | 0.44210122 | 0.61347053 |

7.4.2. OVERALL DISCRIMINATION

Samples of seven cytospecies from the area west of Togo were available for analysis, from four countries, Côte d'Ivoire, Guinea, Mali, and Sierra Leone (see Appendix one).

For this analysis, the single *S. damnosum* s.s. sample was included with the five *S. sirbanum* samples in a single artificial category designated 'Savanna'. This step was justified because the WHO Onchocerciasis Control Programme within this area regards both species to be dangerous vectors to be controlled wherever they are found, and morphometrically they are not distinctive (section 7.4.1.1). The single sample of *S. soubrense* 'B' was included within the rest of *S. soubrense*, because morphometrically this species is not distinctive enough to justify the extra cost involved in estimating more parameters (see Chapter six, and section 7.4.1.11), despite the chromosomal distinctiveness of this taxon (Chapter three, Post 1986).

The 25 character set (excluding wing tuft colouration and abdominal setal colouration) resulted in the matrix of Mahalanobis' squared distances shown as Table 7.117. This matrix shows some unexpected features, for example, while *S. yahense* is chromosomally close to *S. squamosum* (Vajime and Dunbar 1975), it is morphometrically closer to *S. soubrense* and *S. sanctipauli* than it is to *S. squamosum*. All of these distances are significant at $p < 0.001$. The number of flies misclassified using resubstitution was 74 ($e_{res} = 0.1213$, Table 7.118). The data set was too large to obtain an estimate of error rate using the 'leave-one-out' method of Lachenbruch and Mickey (1968), but the relatively large sample sizes probably means that the resubstituted error rate is not seriously biased.

Examining the table of resubstitution (Table 7.119) shows that the species with by far the greatest number of misidentifications is *S. soubrense*, with over 25% of this species being misallocated into other species.

A stepwise discriminant analysis on the 25 character set resulted in the rejection of only four characters. Therefore, the method described in section 7.2.2 was applied until an eleven character subset was obtained:

[V4,V6,V9,V11,V16,V17,V19,V20,V22,V28,V29]

i.e. one thorax, one head, two antennal, four wing and three leg characters.

The matrix of Mahalanobis' squared distances between species using this 11 character subset is shown as Table 7.119. This matrix shows essentially the same features as the full character set. The number of flies misallocated using this subset and resubstitution was 98 ($e_{res}=0.1607$, Table 7.120), while the 'leave-one-out' method resulted in 110 misclassifications ($e_c=0.1803$).

Allocating the flies using the typicality probability approach described in section 7.2.4, with a fly being defined as atypical if the none of the probabilities associated with the distance from each species' mean was greater than 0.01, then 479 (78.52%) were allocated correctly, 14 were atypical (2.29%), 38 (6.23%) were overlapping and 79 (12.95%) were incorrectly allocated. When the flies which were on the overlap of two species were allocated using the appropriate pairwise discriminant statistics, the number of correctly allocated flies rose to 502 (82.295%), the number overlapping fell to 11 (1.8%) and the number wrongly allocated rose to 83 (13.61%).

The null hypotheses of equal wing tuft colouration and equal abdominal setal colouration were both rejected at $p < 0.001$, using a Kruskal-Wallis non-parametric one-way analysis of variance. Two sets of prior probabilities of species membership were calculated, one set from the wing tuft colour alone, the other set from both colour characters, these are shown as Tables 7.121 and 7.122. The method used to calculate these probabilities was described in section 7.2.3. These two sets of priors were calculated because of the extreme influence that abdominal setal colouration had on the analysis; any *S. yahense* flies with abdominal setal colouration category one were automatically classified into another species.

Adjusting the prior probability of a fly's species membership using the priors given in Table 7.121 resulted in 84 flies being wrongly allocated ($e_{\text{res}} = 0.1377$), while adjusting the priors using both colour characters (Table 7.122) resulted in 68 flies being incorrectly allocated ($e_{\text{res}} = 0.1115$).

Table 7.123 gives the standardised canonical variates, showing which of the eleven characters are important in discrimination, also the means of each species on each canonical variate are shown as Table 7.124 to aid interpretation.

The first canonical variate, with a canonical root of 5.2451, is clearly dominated by antennal length 1, such that flies at the positive end of this vector can be expected to have relatively larger antennae than those at the negative end. Thus, *S. sanctipauli* is at the positive end of the vector and 'Savanna' at the negative end. This confirms previous work on the morphology of the *S. damnosum* complex (see e.g. Garms 1977). Other characters with some influence

along this vector are basitarsus length 2 and thorax width, the former is positively loading, the latter negative.

The second canonical variate has a canonical root of 0.9791, which is not very large (Campbell 1982). However, two species are quite well discriminated along this vector (*S. sanctipauli* and *S. squamosum*), so the vector is of importance. The most influential character is wing length 3, which has a large positive loading, so that wing size in *S. squamosum* can be expected to be relatively larger than in *S. sanctipauli*. A relationship between the two mid-leg characters, tibia length 2 and basitarsus length 2 is also of importance along this vector, with tibia length 2 loading positively and basitarsus length 2 loading negatively.

The other canonical variates have canonical roots of only 0.3187 and 0.0799 respectively, and so are not of any importance, especially as neither discriminates one species particularly well.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.32,0.31,0.25,0.2,0.32,0.31,0.31,0.32,0.32,0.32,0.32]

and accounted for 80.31% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable, effectively controlling for size variation, the first canonical root fell to 4.035, showing that although size has some influence along this vector, it is not as important as shape variation. The second canonical root fell to 0.9674 showing that size is not important along this vector.

The mean vectors are shown as Table 7.125, the pooled within-species dispersion matrix as Table 7.126 and the linear discriminant functions as Table 7.127.

To conclude, the overall discriminant analysis of western flies has revealed a great deal of multivariate morphological differentiation between the five taxa examined in this area. The statistics derived in this analysis should be of considerable assistance to attempts to understand more fully the relative vectorial importance of the different taxa in this area.

Table 7.117

Matrix of Mahalanobis' distances between species, 25 character set

| | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|-----------------------|---------|-----------------------|---------------------|---------------------|-------------------|
| Savanna | 0.0 | | | | |
| <i>S. sanctipauli</i> | 47.92 | 0.0 | | | |
| <i>S. soubrense</i> | 25.43 | 11.42 | 0.0 | | |
| <i>S. squamosum</i> | 13.89 | 36.66 | 12.69 | 0.0 | |
| <i>S. yahense</i> | 35.08 | 12.43 | 5.76 | 14.9 | 0.0 |

Table 7.118

Table of re-classifications, using resubstitution, 25 character set

| | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|-----------------------|---------|-----------------------|---------------------|---------------------|-------------------|
| Savanna | 202 | 0 | 1 | 3 | 0 |
| <i>S. sanctipauli</i> | 0 | 32 | 0 | 0 | 3 |
| <i>S. soubrense</i> | 5 | 10 | 139 | 8 | 24 |
| <i>S. squamosum</i> | 2 | 0 | 1 | 84 | 0 |
| <i>S. yahense</i> | 0 | 7 | 10 | 0 | 79 |

$$e_{res} = 74/610 = 0.1213$$

Table 7.119

Prior probability of species membership for each wing tuft colour category

| Wing Tuft Colour | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|------------------|---------|-----------------------|---------------------|---------------------|-------------------|
| 1 | 0.664 | 0.0005 | 0.0476 | 0.2854 | 0.0 |
| 2 | 0.2568 | 0.0156 | 0.2493 | 0.4666 | 0.0027 |
| 3 | 0.0379 | 0.1655 | 0.4659 | 0.2724 | 0.0584 |
| 4 | 0.0013 | 0.4347 | 0.2152 | 0.0393 | 0.3095 |
| 5 | 0.0 | 0.3955 | 0.0344 | 0.002 | 0.5681 |

Table 7.120

Prior probability of species membership for each wing tuft and abdominal setal colour category

| A ¹ | B ² | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|----------------|----------------|---------|-----------------------|---------------------|---------------------|-------------------|
| 1 | 1 | 0.6665 | 0.0006 | 0.0476 | 0.2853 | 0.0 |
| 1 | 2 | 0.2662 | 0.0167 | 0.2497 | 0.4674 | 0.0 |
| 1 | 3 | 0.0399 | 0.1813 | 0.4915 | 0.2873 | 0.0 |
| 1 | 4 | 0.0019 | 0.6312 | 0.3103 | 0.0566 | 0.0 |
| 1 | 5 | 0.0 | 0.9139 | 0.0814 | 0.0046 | 0.0 |
| 2 | 1-5 | 0.0 | 0.0 | 0.0 | 0.0 | 1.0 |

A¹ Abdominal setal colour.

B² Wing tuft colour.

Table 7.121

Matrix of Mahalanobis' distances between species, 11 character subset

| | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|-----------------------|---------|-----------------------|---------------------|---------------------|-------------------|
| Savanna | 0.0 | | | | |
| <i>S. sanctipauli</i> | 43.46 | 0.0 | | | |
| <i>S. soubrense</i> | 20.23 | 9.5 | 0.0 | | |
| <i>S. squamosum</i> | 12.27 | 32.68 | 9.45 | 0.0 | |
| <i>S. yahense</i> | 30.73 | 9.65 | 4.35 | 12.41 | 0.0 |

Table 7.122

Table of re-classifications, using resubstitution (and 'leave-one-out'), 11 character subset

| | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|-----------------------|----------|-----------------------|---------------------|---------------------|-------------------|
| Savanna | 200(197) | 0(0) | 2(2) | 4(7) | 0(0) |
| <i>S. sanctipauli</i> | 0(0) | 34(31) | 1(3) | 0(0) | 0(1) |
| <i>S. soubrense</i> | 5(5) | 15(15) | 121(119) | 13(15) | 32(32) |
| <i>S. squamosum</i> | 4(4) | 0(0) | 4(4) | 78(77) | 1(2) |
| <i>S. yahense</i> | 0(0) | 4(5) | 13(14) | 0(1) | 79(76) |

$$e_{res} = 98/610 = 0.1607$$

$$e_c = 110/610 = 0.1803$$

Table 7.123

Standardised Canonical Variates

| Character | CV I | CV II | CV III | CV IV |
|---------------------|---------|---------|---------|---------|
| Thorax Width | -0.7428 | -0.8216 | -0.0093 | -1.2503 |
| Head Width | -0.3497 | -0.1707 | -0.4332 | 0.3250 |
| Antennal Length 1 | 2.1559 | -0.5048 | -0.7147 | -0.8025 |
| Antennal Segment 4 | 0.3230 | -0.7823 | 0.5356 | 0.2273 |
| Wing Length 1 | -0.3358 | -0.9190 | -0.0470 | 1.6083 |
| Wing Length 2 | -0.5758 | 0.6049 | 0.4459 | -0.6043 |
| Wing Width 2 | -0.0215 | -0.5004 | 1.0569 | 0.8070 |
| Wing Length 3 | 0.5966 | 2.1948 | -0.5248 | -2.0623 |
| Femur Length 1 | 0.2295 | 0.5962 | -2.3321 | 2.0908 |
| Tibia Length 2 | -0.2896 | -1.5068 | 0.6115 | -1.5701 |
| Basitarsus Length 2 | 0.9069 | 1.5298 | 1.8634 | 1.0686 |

Table 7.124

Species Means on Canonical Variates

| Species | CV I | CV II | CV III | CV IV |
|-----------------------|---------|---------|---------|---------|
| Savanna | -2.9028 | -0.4728 | 0.1692 | -0.0403 |
| <i>S. sanctipauli</i> | 3.4005 | -2.1675 | 0.4904 | 0.8259 |
| <i>S. soubrense</i> | 1.5073 | -0.2749 | -0.6867 | -0.1472 |
| <i>S. squamosum</i> | -0.4572 | 1.9737 | -0.1852 | 0.3831 |
| <i>S. yahense</i> | 2.4832 | 0.5487 | 0.9564 | -0.2766 |

Table 7.125

Mean Vectors

| Character | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> |
|---------------------|---------------|-----------------------|---------------------|
| Thorax Width | 859.40403204 | 906.40608000 | 873.10707742 |
| Head Width | 775.71054175 | 831.90817714 | 805.30381935 |
| Antennal Length 1 | 253.08349515 | 336.06857143 | 307.02016129 |
| Antennal Segment 4 | 35.08718447 | 49.91885714 | 42.24000000 |
| Wing Length 1 | 705.99611650 | 771.73714286 | 733.50322581 |
| Wing Length 2 | 435.87378641 | 457.41942857 | 443.85806452 |
| Wing Width 2 | 1383.34075243 | 1488.56614286 | 1412.69195699 |
| Wing Length 3 | 1407.11145631 | 1512.93442857 | 1492.00408602 |
| Femur Length 1 | 597.99495146 | 657.87428571 | 635.50000000 |
| Tibia Length 2 | 597.82776699 | 644.66057143 | 619.17935484 |
| Basitarsus Length 2 | 305.07087379 | 344.64857143 | 326.74032258 |

Table 7.125 (continued)

Mean Vectors

| Character | <i>S. squamosum</i> | <i>S. yahense</i> |
|---------------------|---------------------|-------------------|
| Thorax Width | 855.21798621 | 906.62228125 |
| Head Width | 795.31917241 | 828.68378125 |
| Antennal Length 1 | 276.18275862 | 319.98281250 |
| Antennal Segment 4 | 36.13103448 | 45.00166667 |
| Wing Length 1 | 734.04137931 | 770.69750000 |
| Wing Length 2 | 450.71724138 | 471.03875000 |
| Wing Width 2 | 1411.51097701 | 1495.92739583 |
| Wing Length 3 | 1506.67091954 | 1576.95807292 |
| Femur Length 1 | 631.34344828 | 657.64000000 |
| Tibia Length 2 | 612.96413793 | 648.41500000 |
| Basitarsus Length 2 | 329.43793103 | 352.44218750 |

Table 7.126

Pooled within-species dispersion matrix

| Character | V4 | V6 | V9 | V11 |
|-----------|---------------|---------------|---------------|--------------|
| V4 | 3938.81886706 | 2547.86417923 | 700.67149127 | 122.06071806 |
| V6 | 2547.86417923 | 2170.96003419 | 537.16930604 | 88.17262381 |
| V9 | 700.67149127 | 537.16930604 | 292.41436279 | 37.76227516 |
| V11 | 122.06071806 | 88.17262381 | 37.76227516 | 14.47383944 |
| V16 | 2830.31111161 | 2016.41932163 | 590.33707686 | 94.59387621 |
| V17 | 1707.85388944 | 1227.64588087 | 365.56136579 | 56.38481987 |
| V19 | 4660.59315723 | 3289.08156046 | 933.90669487 | 150.50287360 |
| V20 | 4978.36830465 | 3572.66929859 | 1008.58724661 | 169.08697429 |
| V22 | 2354.74558447 | 1703.22821304 | 485.32662599 | 80.14613543 |
| V28 | 2274.59173427 | 1622.90781954 | 451.03237589 | 74.96679100 |
| V29 | 1208.31487019 | 882.61668440 | 253.18259454 | 43.06575188 |

| Character | V16 | V17 | V19 | V20 |
|-----------|---------------|---------------|---------------|---------------|
| V4 | 2830.31111161 | 1707.85388944 | 4660.59315723 | 4978.36830465 |
| V6 | 2016.41932163 | 1227.64588087 | 3289.08156046 | 3572.66929859 |
| V9 | 590.33707686 | 365.56136579 | 933.90669487 | 1008.58724661 |
| V11 | 94.59387621 | 56.38481987 | 150.50287360 | 169.08697429 |
| V16 | 2734.16453406 | 1507.75055961 | 3817.39933797 | 4281.72764426 |
| V17 | 1507.75055961 | 1044.00879161 | 2383.22635347 | 2514.37284250 |
| V19 | 3817.39933797 | 2383.22635347 | 7629.48029291 | 6894.47454785 |
| V20 | 4281.72764426 | 2514.37284250 | 6894.47454785 | 8168.03898552 |
| V22 | 1973.86763070 | 1181.72550851 | 3123.83941523 | 3457.55908978 |
| V28 | 1910.00548457 | 1149.80753341 | 3058.94929888 | 3339.27956167 |
| V29 | 1035.06316022 | 613.92763037 | 1619.15046300 | 1803.27475551 |

| Character | V22 | V28 | V29 |
|-----------|---------------|---------------|---------------|
| V4 | 2354.74558447 | 2274.59173427 | 1208.31487019 |
| V6 | 1703.22821304 | 1622.90781954 | 882.61668440 |
| V9 | 485.32662599 | 451.03237589 | 253.18259454 |
| V11 | 80.14613543 | 74.96679100 | 43.06575188 |
| V16 | 1973.86763070 | 1910.00548457 | 1035.06316022 |
| V17 | 1181.72550851 | 1149.80753341 | 613.92763037 |
| V19 | 3123.83941523 | 3058.94929888 | 1619.15046300 |
| V20 | 3457.55908978 | 3339.27956167 | 1803.27475551 |
| V22 | 1753.43979004 | 1600.05442471 | 842.56492900 |
| V28 | 1600.05442471 | 1620.39537057 | 819.79925300 |
| V29 | 842.56492900 | 819.79925300 | 486.25898631 |

Table 7.127

Linear Discriminant functions

| | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|----------|--------------|-----------------------|---------------------|---------------------|-------------------|
| CONSTANT | -182.5282093 | -254.03615624 | -224.90401261 | -205.64371422 | -244.71672460 |
| V4 | -0.15031241 | -0.21733740 | -0.20081153 | -0.21698026 | -0.22003282 |
| V6 | 0.34089754 | 0.30583341 | 0.31636906 | 0.32147748 | 0.29192484 |
| V9 | 0.47002013 | 0.87270319 | 0.77099716 | 0.58770168 | 0.78939291 |
| V11 | -0.11107483 | 0.51258684 | 0.02045405 | -0.31465100 | 0.10764204 |
| V16 | -0.09101005 | -0.07669257 | -0.12237300 | -0.13247048 | -0.14632048 |
| V17 | -0.12683996 | -0.27332493 | -0.20646178 | -0.13679257 | -0.18471890 |
| V19 | 0.11571448 | 0.13393345 | 0.10339051 | 0.10209010 | 0.11586675 |
| V20 | 0.12504147 | 0.10727173 | 0.15971974 | 0.18206171 | 0.17637932 |
| V22 | -0.19746609 | -0.16590277 | -0.13631686 | -0.11886083 | -0.20767140 |
| V28 | 0.21674973 | 0.20685222 | 0.17287595 | 0.09688152 | 0.16570609 |
| V29 | -0.37117770 | -0.20279346 | -0.27723383 | -0.16278160 | -0.09361895 |

7.4.3 DISCUSSION

The discriminant analyses presented in this section have revealed that there is considerable multivariate morphometric variation within the *S. damnosum* complex in the area west of the Volta lake.

The two 'savanna' species, *S. damnosum s.s.* and *S. sirbanum* are phenetically similar, with over 10% phenetic overlap, and a considerable proportion of the differences between them being due to size variation. However, the pooled taxon 'savanna' shows very little morphological overlap with any of the other species in this area, the largest being with *S. squamosum*, at about 5%. The differentiation of the 'savanna' flies from the other species involves the relative length of the antenna, and also the relative length of the basitarsal segment of the mid-leg. Size also influences the differentiation of 'savanna' from the other species. but the size-free canonical roots are also the highest when this species is involved than for most other analyses. The ability to identify adult females of this species is of considerable practical importance considering the dangerous vectorial role of *S. damnosum s.s.* and *S. sirbanum*.

Simulium squamosum is also a well isolated species morphometrically, with the maximum overlap being about 5%. This rate of correct identification compares with that using enzyme electrophoresis (Meredith and Townson 1981, Garms and Zillman 1984), so that the two methods together should provide unequivocal identification of members of this species. This contrasts with larval cytotaxonomy, as *S. squamosum* and *S. yahense* are currently only distinguishable as samples in parts of the western area (Surtees and Post unpublished data).

Members of the *S. sanctipauli* subcomplex are morphologically very similar, and *S. yahense* is also very close to these species. Therefore it appears that the chromosomal evolution of the *S. sanctipauli* subcomplex and *S. yahense* has not been paralleled in morphology, or that these species have converged to a common morphology. *Simulium yahense* can be distinguished from the *S. sanctipauli* subcomplex with nearly 96% accuracy using the colour of the abdominal setae, although scoring this character is less objective, and hence more prone to error, than taking measurements.

Table 7.128 summaries the performance of the different allocation methods used in the overall analysis. The method which resulted in the largest proportion of correct identifications was the forced allocation method with prior probabilities adjusted according to both wing tuft colour and abdominal setal colour. The typicality probability method resulted in over 6% of flies lying on an overlap of two or more species. This was a smaller proportion than the equivalent analysis in Togo and Benin because of the larger sample sizes giving narrower approximate confidence intervals. The proportion overlapping fell once the flies had been allocated using typicality probabilities calculated from the appropriate species-pair statistics.

To conclude, there is considerable morphological differentiation between members of the *S. damnosum* complex in the western area. In particular, 'savanna' and *S. squamosum* can be very successfully identified, as can *S. yahense* if abdominal setal colour is used. Members of the *S. sanctipauli* subcomplex cannot be very well distinguished, although the practical importance of this may not be great.

Table 7.128

Comparison of five methods of allocation for the Western area

| | Forced ¹ | Forced ² | Forced ³ | Typicality | Typicality |
|-------------|---------------------|---------------------|---------------------|------------|------------|
| Correct | 512 | 526 | 542 | 479 | 502 |
| Incorrect | 98 | 84 | 68 | 79 | 83 |
| Overlapping | na | na | na | 38 | 11 |
| Atypical | na | na | na | 14 | 14 |

¹Forced allocation without adjusted priors.

²Forced allocation with prior probabilities adjusted for wing tuft colour

³Forced allocation with prior probabilities adjusted for wing tuft colour and abdominal setal colour

⁴Typicality probability without subsequent species pair allocation of overlapping flies

⁵Typicality probability with subsequent species pair allocation of overlapping flies

CHAPTER EIGHT: GLOBAL DISCRIMINATION OF SPECIES.

8.1 INTRODUCTION

Chapter seven has developed the statistics for allocation assuming prior knowledge of the geographic origin of the fly to be allocated, and hence of the possible range of reference species. The purpose of this chapter is to develop the statistics necessary for allocation without prior geographic knowledge.

For this analysis *S. soubrense* and *S. soubrense* 'B' were treated as a single category, *S. soubrense* for the reasons given in section 7.1, as were *S. damnosum* and *S. sirbanum*, which were pooled into the category 'savanna'. Thus there were five reference groups for allocation, *S. soubrense*, *S. sanctipauli*, *S. squamosum*, *S. yahense*, and 'savanna'.

8.2 MATERIALS AND METHODS

The full list of samples is given in Appendix one, and the methods used were the same as described in Chapter seven.

8.3 RESULTS AND DISCUSSION

8.3.1 PAIRWISE DISCRIMINATION

8.3.1.1 Discrimination of 'savanna' and S. sanctipauli

A total of 249 'savanna' flies from eight samples, three *S. damnosum* s.s., and three *S. sirbanum*, from five West African countries, and 61 *S. sanctipauli*, from two samples in two West African countries were examined (Appendix one).

The 25 character set described in Chapter four resulted in a Mahalanobis' squared distance of 55.28 (unbiased $D^2=50.61$, $e_{act}<0.001$), with no misallocations using resubstitution, $e_{res}=0.0$, showing that there is considerable morphological divergence between these species.

A 12 character set derived using stepwise discriminant analysis resulted in a squared distance of 53.69 (unbiased $D^2=51.43$, $e_{act}<0.001$) with no misallocated flies. This 12 character subset was further reduced using the method described in section 7.2.2 to the four character set:

[V4,V9,V13,V29]

i.e. one thorax, two antennal and one leg measurement (Chapter four), which gave a Mahalanobis' squared distance of 44.29 (unbiased $D^2=43.57$, $e_{act}=0.0005$) between the species, and one fly misallocated using resubstitution (a 'savanna' fly misidentified as *S. sanctipauli*), $e_{res}=0.003$.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in one misallocated fly (0.323%), four atypical flies (1.29%) and 305 correctly allocated flies (98.39%).

The null hypothesis of equal wing tuft colour was rejected at $p < 0.001$ (Wilcoxon two sample rank-sum test), so the adjusted prior probabilities of species membership according to a fly's wing tuft colour are given as Table 8.1. This shows that *S. sanctipauli* flies have darker wing tufts than 'savanna'. The effect of adjusting the prior probabilities of species membership according to wing tuft colour was to reduce the resubstituted error rate to zero.

The null hypothesis of equal dispersion was not rejected at $p < 0.001$ using the likelihood ratio test, legitimising the pooling of the individual species' dispersion matrices.

The standardised canonical variate (Table 8.2) shows that three of the characters: the two antennal characters and the leg character have positive loadings on the vector, while thorax width has a negative loading. Thus, *S. sanctipauli*, which is at the positive side of this vector has a relatively larger antenna than 'savanna'.

The first eigenvector of the pooled within-species correlation matrix was a size vector with coefficients:

$$[0.55, 0.48, 0.41, 0.55]$$

and accounted for 64.5% of pooled within-species variance. When the scores along this vector were introduced into the model as a covariable, the canonical root fell from 7.0458 to 4.3566, indicating that size has some influence on discrimination, but that shape differences are important.

The mean vectors are shown as Table 8.3, the pooled within-species dispersion matrix as Table 8.4 and the linear discriminant functions as Table 8.5.

To conclude, there is significant multivariate morphometric differentiation between 'savanna' and *S. sanctipauli*. This variation

includes both size and shape components, but discrimination is still very successful when size variation is controlled for. The four character subset can be expected to classify flies to their correct species in over 98% of cases when it is known *a priori* that just this species pair can be expected.

Adjusting the prior probabilities of species membership according to a fly's wing tuft colouration improves allocation rate, so it is recommended that this should be done for flies of doubtful affinity following typicality probability allocation.

Table 8.1

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------|-----------------------|
| | 'Savanna' | <i>S. sanctipauli</i> |
| 1 | 1.0000 | 0.0000 |
| 2 | 0.9998 | 0.0002 |
| 3 | 0.0012 | 0.9988 |
| 4 | 0.0000 | 1.0000 |
| 5 | 0.0000 | 1.0000 |

Table 8.2

Standardised Canonical Variate for 'Savanna' and *S. sanctipauli*

| Character | Canonical Variate |
|---------------------|-------------------|
| Thorax Width | -1.0102 |
| Antennal Length 1 | 1.5531 |
| Antennal Segment 6 | 1.1886 |
| Basitarsus Length 2 | 0.7587 |

Table 8.3

Mean Vectors for species 'Savanna' and *S. sanctipauli*

| | 'Savanna' | <i>S. sanctipauli</i> |
|---------------------|--------------|-----------------------|
| Thorax Width | 867.74462410 | 909.84725902 |
| Antennal Length 1 | 254.11867470 | 328.20737705 |
| Antennal Segment 6 | 36.08449799 | 50.53508197 |
| Basitarsus Length 2 | 307.67710843 | 347.06803279 |

Table 8.4

Pooled within-species dispersion matrix for species 'Savanna' and *S. sanctipauli*

| Character | V4 | V9 | V13 | V29 |
|-----------|---------------|--------------|-------------|---------------|
| V4 | 3793.50722634 | 451.12961410 | 71.96460652 | 1069.40693381 |
| V9 | 451.12961410 | 212.63157417 | 21.20256590 | 145.85851731 |
| V13 | 71.96460652 | 21.20256590 | 9.01493600 | 22.06170470 |
| V29 | 1069.40693381 | 145.85851731 | 22.06170470 | 399.42636260 |

Table 8.5

Linear Discriminant functions for species 'Savanna' and *S. sanctipauli*

| | 'Savanna' | <i>S. sanctipauli</i> |
|----------|---------------|-----------------------|
| CONSTANT | -185.82239913 | -301.48495611 |
| V4 | -0.02075579 | -0.12623403 |
| V9 | 0.80451694 | 1.11869762 |
| V13 | 1.12638086 | 2.34553278 |
| V29 | 0.46986876 | 0.66882302 |

8.3.1.2 Discrimination of 'savanna' and *S. soubrense*

Two hundred and twenty-two *S. soubrense* from eight samples taken in three West African countries were examined in relation to the 249 'savanna' flies (Appendix one).

The 25 character set resulted in a Mahalanobis' squared distance between species of 20.15 (unbiased $D^2=19.036$, $e_{act}=0.0146$) with nine individuals misallocated using resubstitution, $e_{res}=0.019$. Stepwise

discriminant analysis resulted in an initial subset of 13 characters with a D^2 of 19.63 (unbiased $D^2=19.044$, $e_{act}=0.0146$) and ten misallocated flies ($e_{res}=0.021$). Applying the method described in section 7.2.2 a five character resulted:

[V9,V13,V17,V19,V22]

i.e two antennal, two wing and one leg character, which had a Mahalanobis' squared distance of 17.09 between species (unbiased $D^2=16.87$, $e_{res}=0.02$) and 11 misallocated flies using resubstitution ($e_{res}=0.023$), these being three 'savanna' flies and eight *S. soubrense* misallocated.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in 11 misallocated flies (2.34%), five atypical flies (1.06%), three overlapping flies (0.64%) and 452 correctly allocated flies (95.97%).

The null hypothesis of equal wing tuft colour was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test, so the prior probabilities of species membership according to a fly's wing tuft colour were calculated (Table 8.6). This shows that 'savanna' flies have lighter wing tufts than *S. soubrense*. When the prior probabilities were adjusted, however, the resubstituted error rate rose to 15/471 ($e_{res}=0.038$), due to the pale wing tufted *S. soubrense* from sample V1=19 (Chapter six) being heavily penalised against 'own-group' membership. Including wing tuft colouration in the linear discriminant function resulted in eleven misallocations, nine *S. soubrense* into 'savanna', for the same reasons. Therefore, for 'global' discrimination of these species it is not recommended that wing tuft colouration be used for adjusting prior probabilities.

The likelihood ratio test for equality of dispersion was rejected at $p < 0.0001$, bringing into question the validity of pooling the dispersion matrices. However, the pooled dispersion matrix was used for the practical reasons discussed in section 7.2.5.

The standardised canonical variate (Table 8.7) shows that the most important character in discriminating these two species is antennal length 1. The antennal segment and the femur length of the first leg also load positively on the variable, whilst the wing characters both load negatively. Thus, the relative size of the antenna, which is larger in *S. soubrense*, is the main feature discriminating between this species pair.

The first eigenvector of the pooled within-species correlation matrix was a size vector, had coefficients:

[0.43, 0.34, 0.48, 0.48, 0.49]

and accounted for 71% of pooled within-species variation. When the scores along this vector were introduced as a covariable into the model, the canonical root fell from 4.2758 to 3.8441, indicating that size was not an important discriminant relative to shape differences between these species.

The mean vectors are shown as Table 8.8, the linear discriminant function as Table 8.9 and the pooled within-species dispersion matrix as Table 8.10

To conclude, there is significant multivariate morphometric differentiation between 'savanna' and *S. soubrense*. This variation is mostly shape variation involving relative antennal size. The five character subset can be expected to classify flies to their correct species in over 95% of cases when it is known *a priori* that just this species pair is expected.

Adjusting the prior probabilities of species membership according to a fly's wing tuft colouration is detrimental to error rate, so its use is not recommended.

Table 8.6

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------|---------------------|
| | 'Savanna' | <i>S. soubrense</i> |
| 1 | 0.9181 | 0.0819 |
| 2 | 0.5336 | 0.4664 |
| 3 | 0.1045 | 0.8955 |
| 4 | 0.0118 | 0.9882 |
| 5 | 0.0012 | 0.9988 |

Table 8.7

Standardised Canonical Variate for 'Savanna' and *S. soubrense*

| Character | Canonical Variate |
|--------------------|-------------------|
| Antennal Length 1 | 1.6258 |
| Antennal Segment 6 | 0.7682 |
| Wing Length 2 | -0.6028 |
| Wing Width 2 | -0.5660 |
| Femur Length 1 | 0.4483 |

Table 8.8

Mean Vectors for species 'Savanna' and *S. soubrense*

| | 'Savanna' | <i>S. soubrense</i> |
|--------------------|---------------|---------------------|
| Antennal Length 1 | 254.11867470 | 304.93783784 |
| Antennal Segment 6 | 36.08449799 | 45.50576577 |
| Wing Length 2 | 439.08530120 | 443.02162162 |
| Wing Width 2 | 1389.53391566 | 1401.13479279 |
| Femur Length 1 | 602.69012048 | 633.77135135 |

Table 8.9

Pooled within-species dispersion matrix for species 'Savanna' and *S. soubrense*

| Character | V9 | V13 | V17 |
|-----------|---------------|---------------|---------------|
| V9 | 289.78578427 | 33.96758774 | 340.38985869 |
| V13 | 33.96758774 | 11.24929548 | 44.42128441 |
| V17 | 340.38985869 | 44.42128441 | 1055.46127339 |
| V19 | 928.19454943 | 130.29886217 | 2433.44880791 |
| V22 | 435.82477951 | 59.87858877 | 1141.13146814 |
| Character | V19 | V22 | |
| V9 | 928.19454943 | 435.82477951 | |
| V13 | 130.29886217 | 59.87858877 | |
| V17 | 2433.44880791 | 1141.13146814 | |
| V19 | 7989.75793638 | 3045.75962135 | |
| V22 | 3045.75962135 | 1610.71537461 | |

Table 8.10

Linear Discriminant functions for species 'Savanna' and *S. soubrense*

| | 'Savanna' | <i>S. soubrense</i> |
|----------|---------------|---------------------|
| CONSTANT | -148.26573187 | -188.51546669 |
| V9 | 0.44074801 | 0.66063765 |
| V13 | 0.48058134 | 1.03014963 |
| V17 | -0.12677299 | -0.20340642 |
| V19 | 0.10345703 | 0.07731114 |
| V22 | 0.13123615 | 0.17433713 |

8.3.1.3 Discrimination of 'savanna' and *S. squamosum*

The 25 character set resulted in a Mahalanobis' squared distance of 15.09 between species (unbiased $D^2=14.15$, $e_{act}=0.03$) with 13 misallocated flies ($e_{res}=0.031$). Stepwise discriminant analysis produced an initial subset of 14 characters with a squared distance of 14.78 (unbiased $D^2=14.25$, $e_{act}=0.0295$) and 15 flies misallocated, $e_{res}=0.036$. Applying the dimension reduction technique described in section 7.2.2 resulted in a six character subset:

[V4,V9,V12,V20,V28,V29]

i.e. one thorax, two antennal, one wing and two leg characters, giving a Mahalanobis' squared distance between species of 11.85 (unbiased $D^2=11.65$, $e_{act}=0.044$) and 19 flies misallocated ($e_{res}=0.045$), ten 'savanna' flies into *S. squamosum*, nine *S. squamosum* flies into 'savanna'.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in 13 misallocated flies (3.09%), three atypical flies (0.71%) six overlapping flies (1.43%) and 399 correctly allocated flies (94.77%).

The null hypothesis of equal wing tuft colour was rejected at $p<0.001$ using a Wilcoxon two sample rank sum test. The adjusted prior probabilities of species membership were therefore calculated and are shown as Table 8.11. When the prior probabilities were adjusted according to a fly's wing tuft colour, the number misallocated remained 19, as it did when wing tuft colouration was included in the linear discriminant function. Therefore no extra discriminatory power was provided by including this character for discriminating between these species.

The null hypothesis of equal dispersion was not rejected at $p=0.0001$, legitimising the use of the pooled within-species dispersion matrix (see section 7.2.5)

The standardised canonical variate (Table 8.12) shows the most important characters in discrimination are the length of the basitarsus of the second leg, length of the tibia of the second leg, wing length and thorax width. Antennal measurements have only a minor impact on discrimination. Wing length and basitarsal length load positively, while thorax width and tibia length load negatively.

The first eigenvector of the pooled within-species correlation matrix was a size vector with coefficients:

[0.44, 0.35, 0.30, 0.44, 0.45, 0.45]

and accounted for 76% of pooled-within species variation. When the scores along this vector were introduced into the model as a covariable, the canonical root fell from 2.8767 to 2.0056, showing that size has some influence as a discriminatory factor, but that shape variation is more important than pure size.

The mean vectors are shown as Table 8.13, the pooled within-species dispersion matrix as Table 8.14 and the linear discriminant functions as Table 8.15.

To conclude, there is significant multivariate morphometric differentiation between 'savanna' and *S. squamosum*. This variation is mostly shape variation involving the thorax width, wing length and the mid-leg. The six character subset can be expected to classify flies to their correct species in nearly 95% of cases when it is known *a priori* that just this species pair is expected.

Adjusting the prior probabilities of species membership according to a fly's wing tuft colouration does not improve allocation rate, so it is not recommended that this should be done.

Table 8.11

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------|---------------------|
| | 'Savanna' | <i>S. squamosum</i> |
| 1 | 0.6691 | 0.3309 |
| 2 | 0.2589 | 0.7411 |
| 3 | 0.0569 | 0.9431 |
| 4 | 0.0103 | 0.9897 |
| 5 | 0.0018 | 0.9982 |

Table 8.12

Standardised Canonical Variate for 'Savanna' and *S. squamosum*

| Character | Canonical Variate |
|---------------------|-------------------|
| Thorax Width | -1.5329 |
| Antennal Length 1 | 0.5197 |
| Antennal Segment 5 | 0.2363 |
| Wing Length 3 | 1.2627 |
| Tibia Length 2 | -1.4962 |
| Basitarsus Length 2 | 2.2498 |

Table 8.13

Mean Vectors for species 'Savanna' and *S. squamosum*

| | 'Savanna' | <i>S. squamosum</i> |
|---------------------|---------------|---------------------|
| Thorax Width | 867.74462410 | 891.07350698 |
| Antennal Length 1 | 254.11867470 | 279.50755814 |
| Antennal Segment 5 | 34.29670683 | 37.76953488 |
| Wing Length 3 | 1414.92809237 | 1540.09226744 |
| Tibia Length 2 | 601.62313253 | 637.39744186 |
| Basitarsus Length 2 | 307.67710843 | 343.40406977 |

Table 8.14

Pooled within-species dispersion matrix for species 'Savanna' and *S. squamosum*

| Character | V4 | V9 | V12 |
|-----------|---------------|---------------|---------------|
| V4 | 4488.85965498 | 633.61858811 | 117.91589590 |
| V9 | 633.61858811 | 238.80768696 | 21.98895923 |
| V12 | 117.91589590 | 21.98895923 | 10.80613189 |
| V20 | 5081.60222426 | 853.06093478 | 140.29902677 |
| V28 | 2648.98664488 | 430.14450062 | 75.35378894 |
| V29 | 1413.40302741 | 230.21063521 | 41.07309918 |
| Character | V20 | V28 | V29 |
| V4 | 5081.60222426 | 2648.98664488 | 1413.40302741 |
| V9 | 853.06093478 | 430.14450062 | 230.21063521 |
| V12 | 140.29902677 | 75.35378894 | 41.07309918 |
| V20 | 7492.71565724 | 3441.29744197 | 1837.01409675 |
| V28 | 3441.29744197 | 1879.71363826 | 956.95545023 |
| V29 | 1837.01409675 | 956.95545023 | 553.35364360 |

Table 8.15

Linear Discriminant functions for species 'Savanna' and *S. squamosum*

| | 'Savanna' | <i>S. squamosum</i> |
|----------|---------------|---------------------|
| CONSTANT | -183.65011147 | -225.40730245 |
| V4 | -0.06024785 | -0.13796231 |
| V9 | 0.71732731 | 0.80740560 |
| V12 | 1.02865257 | 1.24842656 |
| V20 | 0.26325429 | 0.30419872 |
| V28 | -0.06744587 | -0.17762073 |
| V29 | -0.42217823 | -0.15829280 |

8.3.1.4 Discrimination of 'savanna' and *S. yahense*

The 25 character set resulted in a Mahalanobis' squared distance between species of 44.28 (unbiased $D^2=40.92$, $e_{act}=0.0007$) with no flies misallocated using resubstitution ($e_{res}=0.0$). Stepwise discriminant analysis reduced this to a fourteen character subset with a D^2 of 43.58 (unbiased $D^2=41.674$, $e_{act}=0.0006$) and a single fly misallocated ($e_{res}=0.003$). Applying the method for dimension re-

duction described in section 7.2.2, a four character subset was derived,

[V4,V9,V20,V29]

i.e., one thorax, one antennal, one wing and one leg character, with a Mahalanobis' squared distance of 29.6 between species (unbiased $D^2 = 29.17, e_{act} = 0.0035$) and a single fly misallocated as *S. yahense* ($e_{res} = 0.003$). This fly had a posterior probability of species membership (0.4999,0.5001) for 'savanna' and *S. yahense* respectively, and the typicality probability confidence intervals were (0.0118,0.0067) for 'savanna' and (0.0129,0.0062) for *S. yahense*, showing that the fly was in reality atypical of both species at $p < 0.01$.

Allocation using typicality probability of species membership with atypicality defined at $\alpha = 0.01$ resulted in no misallocated flies, six atypical flies (1.74%) one overlapping fly (0.29%) and 338 correctly allocated flies (97.97%).

The null hypotheses of equal wing tuft colouration and equal abdominal setal colouration were both rejected at $p < 0.001$ using a Wilcoxon two-samples rank sum test. The prior probabilities of species membership for each wing tuft category are shown as Table 8.16. No flies were misallocated when the wing tuft colouration was included in the linear discriminant function or when the prior probabilities were adjusted according to wing tuft colouration. By contrast, four *S. yahense* flies were misallocated when prior probabilities were adjusted according to abdominal setal colouration; these flies were *S. yahense* with colour category one (Chapter four). The problems with using the character objectively are discussed in Chapter nine.

The null hypothesis of equal dispersion was rejected at $p < 0.001$ using the likelihood ratio test. However for the practical and statistical reasons discussed in section 7.2.5 the dispersion matrices were pooled.

The standardised canonical variate (Table 8.17) shows that the most important contrast in characters responsible for discrimination is between thorax width and antennal length 1, with the other two characters having only a minor influence on discrimination. *Simulium yahense* is at the positive end of this variable, having a larger antenna relative to 'savanna'.

The first eigenvector of the pooled within-species correlation matrix was a size vector, with coefficients

$$[0.52, 0.43, 0.52, 0.52]$$

and accounted for 82% of pooled within-species variation. When the scores along this vector were introduced as a covariable, controlling for size, the canonical root fell from 5.98 to 3.535, indicating that size has some influence on discrimination, but that shape differences are more important, especially when it is realised that only 18% of pooled within-species variation is involved in discrimination once size is controlled for.

The mean vectors are shown as Table 8.18, the linear discriminant functions as Table 8.19 and the pooled within-species dispersion matrix as Table 8.20.

To conclude, there is significant multivariate morphometric differentiation between 'savanna' and *S. yahense*. This differentiation involves both size and shape variation but in the absence of size there is still considerable discrimination. The shape variation involves a relationship between thorax width and antennal length 1.

The four character subset can be expected to classify flies to their correct species in nearly 98% of cases when it is known *a priori* that just this species pair is expected.

Adjusting the prior probabilities of species membership according to a fly's wing tuft colouration improves allocation rate, so it is recommended that this should be done in cases of doubt following typicality probability allocation. Adjusting the prior probabilities of species membership according to a fly's abdominal setal colouration does not improve allocation rate because of the problem that abdominal setal colouration is not 100% diagnostic for *S. yahense*. Therefore it is not recommended that the character be used routinely but instead it should be used in cases of doubt.

Table 8.16

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------|-------------------|
| | 'Savanna' | <i>S. yahense</i> |
| 1 | 1.0000 | 0.0000 |
| 2 | 1.0000 | 0.0000 |
| 3 | 0.0978 | 0.9022 |
| 4 | 0.0000 | 1.0000 |
| 5 | 0.0000 | 1.0000 |

Table 8.17

Standardised Canonical Variate for 'Savanna' and *S. yahense*

| Character | Canonical Variate |
|---------------------|-------------------|
| Thorax Width | -1.4570 |
| Antennal Length 1 | 2.1645 |
| Wing Length 3 | 0.5252 |
| Basitarsus Length 2 | 0.8300 |

Table 8.18

Mean Vectors for species 'Savanna' and *S. yahense*

| Character | 'Savanna' | <i>S. yahense</i> |
|---------------------|---------------|-------------------|
| Thorax Width | 867.74462410 | 906.62228125 |
| Antennal Length 1 | 254.11867470 | 319.98281250 |
| Wing Length 3 | 1414.92809237 | 1576.95807292 |
| Basitarsus Length 2 | 307.67710843 | 352.44218750 |

Table 8.19

Linear Discriminant functions for species 'Savanna' and *S. yahense*

| | 'Savanna' | <i>S. yahense</i> |
|----------|---------------|-------------------|
| CONSTANT | -178.12879395 | -263.32515247 |
| V4 | -0.10363023 | -0.22159312 |
| V9 | 0.70408309 | 1.05637118 |
| V20 | 0.24719283 | 0.27282018 |
| V29 | -0.26813365 | -0.11546584 |

Table 8.20

Pooled within-species dispersion matrix for species 'Savanna' and *S. yahense*

| Character | V4 | V9 | V20 | V29 |
|-----------|---------------|--------------|---------------|---------------|
| V4 | 4224.07247057 | 635.18361138 | 4820.91533268 | 1243.52743755 |
| V9 | 635.18361138 | 244.49179243 | 834.35947942 | 217.97832669 |
| V20 | 4820.91533268 | 834.35947942 | 7166.92779778 | 1657.94539485 |
| V29 | 1243.52743755 | 217.97832669 | 1657.94539485 | 472.76026577 |

8.3.1.5 Discrimination of Simulium sanctipauli and S. soubrense

The 25 character set resulted in a Mahalanobis' squared distance of 6.81 between species, (unbiased $D^2=6.18$, $e_{act}=0.1069$) with 25 flies misallocated using resubstitution ($e_{res}=0.088$). A stepwise discriminant analysis resulted in an initial 14 character subset with a squared distance of 6.43 (unbiased distance=6.09, $e_{act}=0.1086$), with 27 misallocated flies. The method for dimension reduction given in section 7.2.2 resulted in a nine character subset:

[V6,V14,V15,V16,V17,V19,V20,V24,V27]

i.e. one head, two antennal, four wing and two leg characters, having a Mahalanobis' squared distance of 5.37 (unbiased $D^2=5.18$, $e_{act}=0.1276$) and 29 misallocated flies ($e_{res}=0.1025$), seven of which were misallocated *S. sanctipauli*, 22 were misallocated *S. soubrense*.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in 19 misallocated flies (6.7%), nine atypical flies (3.18%), 25 overlapping flies (8.83%) and 230 correctly allocated flies (81.27%).

The null hypothesis of equal wing tuft colouration was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test. Therefore the prior probabilities of species membership were calculated according to a fly's wing tuft colouration using the method described in Chapter seven; these are shown as Table 8.21. When the prior probabilities were altered in this way 28 flies were misallocated ($e_{res}=0.0989$), while 27 were misallocated when wing tuft colouration was included in the linear discriminant function ($e_{res}=0.095$), so that this character is having only minor influence on discrimination.

The null hypothesis of equal dispersion was rejected using the likelihood ratio test at $p<0.0001$. The dispersion matrices were still

pooled for the practical and statistical reasons given in section 7.2.5.

The standardised canonical variate (Table 8.22) shows that the most important contrast of characters in discrimination of these two species is between wing length 3 and basitarsus length 1. The canonical root was only 0.915 reflecting the relatively poor discrimination between these species.

The first eigenvector of the pooled within-species correlation matrix was mainly a size vector with coefficients:

[0.34,0.21,0.23,0.36,0.36,0.35,0.36,0.37,0.37]

and accounted for 73% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 0.915 to 0.7255, indicating that size variation is a significant proportion of the between-species variation. The importance of size variation in discrimination is discussed in greater detail in Chapter nine.

The mean vectors are shown as Table 8.23, the linear discriminant functions as Table 8.24 and the pooled within-species dispersion matrix as Table 8.25.

To conclude, there is significant multivariate morphometric differentiation between *S. sanctipauli* and *S. soubrense*, although this differentiation is not large. The differentiation involves both size and shape variation but size variation is a considerable component of discrimination. The shape variation involves a relationship between wing length 3 and basitarsus length 1. The nine character subset can be expected to classify flies to their correct species in over 80% of cases when it is known *a priori* that just this species pair is expected.

Adjusting the prior probabilities of species membership according to a fly's wing tuft colouration does not significantly improve allocation rate, so it is not recommended that the character be used routinely.

Table 8.21

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------------------|---------------------|
| | <i>S. sanctipauli</i> | <i>S. soubrense</i> |
| 1 | 0.1453 | 0.8547 |
| 2 | 0.2559 | 0.7441 |
| 3 | 0.4103 | 0.5897 |
| 4 | 0.5846 | 0.4154 |
| 5 | 0.7401 | 0.2599 |

Table 8.22

Standardised Canonical Variate for *S. sanctipauli* and *S. soubrense*

| Character | Canonical Variate |
|---------------------|-------------------|
| Head Width | -0.6145 |
| Antennal Segment 7 | 0.4369 |
| Antennal Segment 8 | 0.3857 |
| Wing Length 1 | 0.8474 |
| Wing Length 2 | -0.9158 |
| Wing Width 2 | 0.7277 |
| Wing Length 3 | -1.7100 |
| Basitarsus Length 1 | 1.0585 |
| Femur Length 2 | 0.7642 |

Table 8.23

Mean Vectors for species *S. sanctipauli* and *S. soubrense*

| Character | <i>S. sanctipauli</i> | <i>S. soubrense</i> |
|---------------------|-----------------------|---------------------|
| Head Width | 834.72223279 | 806.88387027 |
| Antennal Segment 7 | 49.57967213 | 44.30486486 |
| Antennal Segment 8 | 48.38032787 | 43.55639640 |
| Wing Length 1 | 774.45639344 | 730.55351351 |
| Wing Length 2 | 461.91540984 | 443.02162162 |
| Wing Width 2 | 1479.41942623 | 1401.13479279 |
| Wing Length 3 | 1522.32909836 | 1478.64736486 |
| Basitarsus Length 1 | 461.10885246 | 432.29513514 |
| Femur Length 2 | 694.68786885 | 659.01405405 |

Table 8.24

Linear Discriminant functions for species *S. sanctipauli* and *S. soubrense*

| | <i>S. sanctipauli</i> | <i>S. soubrense</i> |
|----------|-----------------------|---------------------|
| CONSTANT | -221.66933357 | -198.01356901 |
| V6 | 0.28345358 | 0.31523435 |
| V14 | 1.68731319 | 1.44047546 |
| V15 | 0.69683414 | 0.47216568 |
| V16 | 0.02301751 | -0.01450255 |
| V17 | -0.24375687 | -0.17865843 |
| V19 | 0.05061038 | 0.03318706 |
| V20 | -0.02458801 | 0.01642916 |
| V24 | -0.12400258 | -0.20545171 |
| V27 | 0.29346838 | 0.25045209 |

Table 8.25

Pooled within-species dispersion matrix for species *S. sanctipauli* and *S. soubrense*

| Character | V6 | V14 | V15 |
|-----------|---------------|---------------|---------------|
| V6 | 1883.76682660 | 62.88650931 | 70.53286981 |
| V14 | 62.88650931 | 12.15178626 | 8.45178754 |
| V15 | 70.53286981 | 8.45178754 | 11.92571356 |
| V16 | 1647.19040184 | 66.53560759 | 77.22545795 |
| V17 | 1054.11490471 | 40.37357351 | 46.69724907 |
| V19 | 2883.48344565 | 127.20979513 | 149.01901390 |
| V20 | 3179.23135670 | 126.00207483 | 155.47490537 |
| V24 | 1001.93177282 | 39.59252467 | 45.24714709 |
| V27 | 1434.61892441 | 51.93177857 | 58.99575191 |
| Character | V16 | V17 | V19 |
| V6 | 1647.19040184 | 1054.11490471 | 2883.48344565 |
| V14 | 66.53560759 | 40.37357351 | 127.20979513 |
| V15 | 77.22545795 | 46.69724907 | 149.01901390 |
| V16 | 2421.69212195 | 1399.57821112 | 3610.91211572 |
| V17 | 1399.57821112 | 1006.13740260 | 2345.59235201 |
| V19 | 3610.91211572 | 2345.59235201 | 8361.24114879 |
| V20 | 4154.26366871 | 2542.01867686 | 7626.77650072 |
| V24 | 1198.99741830 | 764.64662573 | 2027.15060422 |
| V27 | 1686.60200843 | 1070.32927098 | 2889.49389325 |
| Character | V20 | V24 | V27 |
| V6 | 3179.23135670 | 1001.93177282 | 1434.61892441 |
| V14 | 126.00207483 | 39.59252467 | 51.93177857 |
| V15 | 155.47490537 | 45.24714709 | 58.99575191 |
| V16 | 4154.26366871 | 1198.99741830 | 1686.60200843 |
| V17 | 2542.01867686 | 764.64662573 | 1070.32927098 |
| V19 | 7626.77650072 | 2027.15060422 | 2889.49389325 |
| V20 | 9045.06737298 | 2304.48047417 | 3224.88183294 |
| V24 | 2304.48047417 | 769.10953724 | 977.96635349 |
| V27 | 3224.88183294 | 977.96635349 | 1484.68711308 |

8.3.1.6 Discrimination of *Simulium sanctipauli* and *S. squamosum*

The squared Mahalanobis' distance between species using the full 25 character set was 32.18 (unbiased $D^2=28.558$, $e_{act}=0.0038$) with no flies misallocated using resubstitution. A stepwise discriminant analysis produced a 13 character subset which gave a Mahalanobis' squared distance of 30.93 (unbiased $D^2=29.06$, $e_{act}=0.0038$) and no misallocated flies using resubstitution ($e_{res}=0.0$).

The dimension reduction technique given in section 7.2.2 resulted in a six character subset:

[V10,V13,V16,V18,V20,V22]

i.e. two antennal, three wing and one leg character, with a Mahalanobis' squared distance of 25.16 (unbiased $D^2=24.4$, $e_{act}=0.0068$) and no misallocated flies using resubstitution ($e_{res}=0.0$).

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in no misallocated flies, six atypical flies (2.58%), no overlapping flies and 227 correctly allocated flies (97.43%).

The null hypothesis of equal wing tuft colouration was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test, therefore the prior probabilities of species membership for each wing tuft colouration category were calculated using the method described in Chapter seven and are shown as Table 8.26. When the prior probabilities were adjusted no flies were misallocated using resubstitution, as was also the case when wing tuft colouration was included in the linear discriminant function.

The null hypothesis of equal dispersion was rejected at $p<0.001$ using the likelihood ratio test. For the reasons given in section 7.2.5 the pooled dispersion matrix was still used.

The standardised canonical variate (Table 8.27) shows that most discrimination between these species is achieved by antennal length 2 with respect to the other characters. Wing length 3 also has some opposite influence. *Simulium sanctipauli* is at the positive end of this variable indicating that this species has a relatively longer antenna than *S. squamosum*.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.37,0.32,0.43,0.43,0.43,0.44]

and accounted for 79.7% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 4.9 to 4.82 showing that size variation is negligible in discriminating between these species, despite the fact that nearly 80% of within-species variation is size variation.

The mean vectors are shown as Table 8.28, the linear discriminant functions as Table 8.29 and the pooled within-species dispersion matrix as Table 8.30.

To conclude, there is significant multivariate morphometric differentiation between *S. sanctipauli* and *S. squamosum*. The differentiation involves only shape variation which is the relative length of the antenna. The six character subset can be expected to classify flies to their correct species in over 97% of cases when it is known *a priori* that just this species pair is expected.

Adjusting the prior probabilities of species membership according to a fly's wing tuft colouration does not alter an already very good allocation rate, so it is recommended that the character be used only in cases of doubt, following typicality probability allocation.

Table 8.26

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------------------|---------------------|
| | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
| 1 | 0.0013 | 0.9987 |
| 2 | 0.0441 | 0.9559 |
| 3 | 0.6255 | 0.3745 |
| 4 | 0.9837 | 0.0163 |
| 5 | 0.9995 | 0.0005 |

Table 8.27

Standardised Canonical Variate for *S. sanctipauli* and *S. squamosum*

| Character | Canonical Variate |
|--------------------|-------------------|
| Antennal Length 2 | 1.9016 |
| Antennal Segment 6 | 0.7835 |
| Wing Length 1 | 0.2548 |
| Wing Width 1 | -0.3820 |
| Wing Length 3 | -0.7374 |
| Femur Length 1 | -0.3777 |

Table 8.28

Mean Vectors for species *S. sanctipauli* and *S. squamosum*

| Character | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
|--------------------|-----------------------|---------------------|
| Antennal Length 2 | 490.54819672 | 415.62558140 |
| Antennal Segment 6 | 50.53508197 | 40.36488372 |
| Wing Length 1 | 774.45639344 | 758.70976744 |
| Wing Width 1 | 1041.64434426 | 1048.84267442 |
| Wing Length 3 | 1522.32909836 | 1540.09226744 |
| Femur Length 1 | 658.95737705 | 654.56023256 |

Table 8.29

Linear Discriminant functions for species *S. sanctipauli* and *S. squamosum*

| | <i>S. sanctipauli</i> | <i>S. squamosum</i> |
|----------|-----------------------|---------------------|
| CONSTANT | -218.41358418 | -188.35870422 |
| V10 | 0.53891100 | 0.31238504 |
| V13 | 1.16221940 | 0.46913288 |
| V16 | -0.20335053 | -0.22676802 |
| V18 | 0.09262248 | 0.12133480 |
| V20 | 0.24567002 | 0.28775022 |
| V22 | -0.30237415 | -0.26036749 |

Table 8.30

Pooled within-species dispersion matrix for species *S. sanctipauli* and *S. squamosum*

| Character | V10 | V13 | V16 |
|-----------|---------------|---------------|---------------|
| V10 | 686.32956079 | 60.19675399 | 1066.78068583 |
| V13 | 60.19675399 | 12.12667542 | 101.81901954 |
| V16 | 1066.78068583 | 101.81901954 | 2942.82959999 |
| V18 | 1274.19712450 | 127.45080241 | 3162.78558166 |
| V20 | 1724.89487348 | 162.55085422 | 4307.87779348 |
| V22 | 906.58547657 | 89.40313341 | 2221.13883910 |
| Character | V18 | V20 | V22 |
| V10 | 1274.19712450 | 1724.89487348 | 906.58547657 |
| V13 | 127.45080241 | 162.55085422 | 89.40313341 |
| V16 | 3162.78558166 | 4307.87779348 | 2221.13883910 |
| V18 | 4461.52969955 | 5128.11500962 | 2722.01904411 |
| V20 | 5128.11500962 | 7697.39292593 | 3592.05175656 |
| V22 | 2722.01904411 | 3592.05175656 | 2038.62453148 |

8.3.1.7 Discrimination of *Simulium sanctipauli* and *S. yahense*

The squared Mahalanobis' distance between species using the 25 character set was 8.27 (unbiased $D^2=6.88$, $e_{act}=0.095$) with 14 misallocated flies using resubstitution ($e_{res}=0.089$). A stepwise discriminant analysis resulted in an initial subset of 10 characters giving a Mahalanobis' squared distance of 7.33 (unbiased $D^2=6.86$, $e_{act}=0.0952$) and 14 misallocated flies using resubstitution

($e_{res}=0.089$). Applying the method described in section 7.2.2 for dimension reduction in the context of discrimination, a six character subset resulted:

[V3,V11,V16,V17,V20,V28]

i.e. one thorax, one antennal, three wing and one leg measurement. This character subset resulted in a Mahalanobis' squared distance of 6.07 (unbiased $D^2 = 5.796$, $e_{act}=0.1143$) and 14 misallocated flies ($e_{res}=0.089$), seven of each species into the other.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in 9 misallocated flies (5.73%), three atypical flies (1.9%), 11 overlapping flies (7.0%) and 134 correctly allocated flies (85.35%).

The null hypotheses of equal wing tuft colouration and equal abdominal setal colouration were both rejected at $p<0.001$ using Wilcoxon two-sample rank sum tests. The prior probabilities of species membership according to a fly's wing tuft colouration were calculated and are shown as Table 8.31. When the prior probabilities were adjusted for wing tuft colouration the number of flies misallocated fell to eight ($e_{res}=0.051$), with four from each species misallocated. Including wing tuft colouration in the linear discriminant function resulted in 14 misallocated flies ($e_{res}=0.089$), nine *S. yahense* classified as *S. sanctipauli*, and five *S. sanctipauli* classified as *S. yahense*. Including abdominal setal colouration either in the linear discriminant function or by adjusting the prior probabilities resulted in four flies being misallocated ($e_{res}=0.026$), these flies being the four *S. yahense* with abdominal setal colouration character state one. The special nature of this character is discussed in further detail in Chapter nine.

The null hypothesis of equal dispersion was rejected at $p < 0.001$ using the likelihood ratio test. For the reasons given in section 7.2.5 the dispersion matrices were still pooled.

The standardised canonical variate (Table 8.32) shows that the main discrimination between these species is effected by a contrast between wing length 3 and tibia length 2.

The first eigenvector of the pooled within-species correlation matrix was a size vector with coefficients:

[0.39, 0.32, 0.43, 0.43, 0.43, 0.43]

and accounted for 77.3% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 1.4606 to 1.4386 showing that size variation is negligible in discriminating between these species.

The mean vectors are shown as Table 8.33, the linear discriminant functions as Table 8.34 and the pooled within-species dispersion matrix as Table 8.35.

To conclude, there is significant multivariate morphometric differentiation between *S. sanctipauli* and *S. yahense*. This differentiation involves mainly shape variation as in the absence of size there is still quite good discrimination. The shape variation involves a relationship between wing length 3 and femur length 2. The six character subset can be expected to classify flies to their correct species in over 85% of cases when it is known *a priori* that just this species pair is expected.

Adjusting the prior probabilities of species membership according to a fly's wing tuft colouration does not greatly improve allocation rate, so it is not recommended that this should be done. Adjusting the prior probabilities of species membership according to a fly's

abdominal setal colouration improves allocation rate but because of the problem that abdominal setal colouration is not 100% diagnostic for *S. yahense*, it is recommended that the character be used only in cases of doubt following typicality probability allocation.

Table 8.31

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|-----------------------|-------------------|
| | <i>S. sanctipauli</i> | <i>S. yahense</i> |
| 1 | 0.9994 | 0.0006 |
| 2 | 0.9944 | 0.0056 |
| 3 | 0.9526 | 0.0474 |
| 4 | 0.6949 | 0.3051 |
| 5 | 0.2051 | 0.7949 |

Table 8.32

Standardised Canonical Variate for *S. sanctipauli* and *S. yahense*

| Character | Canonical Variate |
|--------------------|-------------------|
| Thorax Length | -0.3837 |
| Antennal Segment 4 | -0.7846 |
| Wing Length 1 | -0.7147 |
| Wing Length 2 | 1.0096 |
| Wing Length 3 | 2.2319 |
| Tibia Length 2 | -1.4841 |

Table 8.33

Mean Vectors for species *S. sanctipauli* and *S. yahense*

| Character | <i>S. sanctipauli</i> | <i>S. yahense</i> |
|--------------------|-----------------------|-------------------|
| Thorax Length | 660.35632131 | 650.18355000 |
| Antennal Segment 4 | 48.33967213 | 45.00166667 |
| Wing Length 1 | 774.45639344 | 770.69750000 |
| Wing Length 2 | 461.91540984 | 471.03875000 |
| Wing Length 3 | 1522.32909836 | 1576.95807292 |
| Tibia Length 2 | 650.97245902 | 648.41500000 |

Table 8.34

Linear Discriminant functions for species *S. sanctipauli* and *S. yahense*

| | <i>S. sanctipauli</i> | <i>S. yahense</i> |
|----------|-----------------------|-------------------|
| CONSTANT | -174.22329721 | -189.77089675 |
| V3 | 0.06924040 | 0.04793217 |
| V11 | -0.06203030 | -0.45841296 |
| V16 | -0.17491616 | -0.20837892 |
| V17 | 0.15615325 | 0.23834071 |
| V20 | 0.20421750 | 0.26670128 |
| V28 | 0.08935998 | -0.00499819 |

Table 8.35

Pooled within-species dispersion matrix for species *S. sanctipauli* and *S. yahense*

| Character | V3 | V11 | V16 |
|-----------|---------------|---------------|---------------|
| V3 | 1955.75824597 | 108.85775937 | 1687.92559940 |
| V11 | 108.85775937 | 21.25674404 | 142.22053115 |
| V16 | 1687.92559940 | 142.22053115 | 2783.37874972 |
| V17 | 940.42740710 | 77.90040328 | 1345.25345219 |
| V20 | 2775.62344008 | 228.90582445 | 3836.82413251 |
| V28 | 1267.40300477 | 90.49433967 | 1786.60875833 |
| Character | V17 | V20 | V28 |
| V3 | 940.42740710 | 2775.62344008 | 1267.40300477 |
| V11 | 77.90040328 | 228.90582445 | 90.49433967 |
| V16 | 1345.25345219 | 3836.82413251 | 1786.60875833 |
| V17 | 901.83462300 | 1994.79639946 | 993.08270767 |
| V20 | 1994.79639946 | 7076.81380738 | 2895.72912168 |
| V28 | 993.08270767 | 2895.72912168 | 1509.68528084 |

8.3.1.8 Discrimination of *Simulium soubrense* and *S. squamosum*

The squared Mahalanobis' distance between species using the 25 character set was 13.15 (unbiased $D^2=12.28$, $e_{act}=0.0399$) with 13 flies misallocated using resubstitution ($e_{res}=0.033$). A stepwise discriminant analysis produced a 12 character subset which gave a Mahalanobis' squared distance of 12.78 (unbiased $D^2=12.356$, $e_{act}=0.0394$) and 15 misallocated flies using resubstitution ($e_{res}=0.038$).

The dimension reduction technique described in Chapter seven resulted in a seven character subset:

[V10, V13, V18, V19, V23, V28, V29]

i.e. two antennal, two wing and three leg characters, with a Mahalanobis' squared distance of 11.56 (unbiased $D^2=11.296$, $e_{act}=0.0464$) and 18 misallocated flies using resubstitution ($e_{res}=0.046$), of which 13 were *S. soubrense* classified as *S. squamosum* and five were *S. squamosum* classified as *S. soubrense*.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in 16 misallocated flies (4.06%), nine atypical flies (2.3%), three overlapping flies (0.76%) and 366 correctly allocated flies (92.89%).

The null hypothesis of equal wing tuft colouration was rejected at $p<0.001$ using a Wilcoxon two-sample rank sum test. The prior probabilities of species membership for each wing tuft colouration category were calculated using the method described in Chapter seven and are shown as Table 8.36. When the prior probabilities were adjusted the number of flies misallocated rose to 22 ($e_{res}=0.056$) whilst including wing tuft colouration in the linear discriminant function resulted in 18 misallocated flies ($e_{res}=0.046$). Therefore wing tuft colouration does not improve discrimination between these species.

The null hypothesis of equal dispersion was rejected at $p<0.001$ using the likelihood ratio test. For the reasons given in section 7.2.5 the pooled dispersion matrix was still used.

The standardised canonical variate (Table 8.37) shows that the main contrast in discriminating these species involves tibia length 1, and ^{basitarsus} length 2 on the positive side, and antennal length 2 and tibia length 2 on the negative side. *Simulium squamosum* is at the positive end of this vector.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

$$[0.36, 0.28, 0.39, 0.38, 0.40, 0.41, 0.40]$$

and accounted for 80.2% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 2.8576 to 2.6199 showing the

negligible effect of size in discriminating between these species, despite the fact 80% of within-species variation is size variation.

The mean vectors are shown as Table 8.38, the linear discriminant functions as Table 8.39 and the pooled within-species dispersion matrix as Table 8.40.

To conclude, there is significant multivariate morphometric differentiation between *S. soubrense* and *S. squamosum*. This differentiation involves mainly shape variation as in the absence of size there is still quite good discrimination. The seven character subset can be expected to classify flies to their correct species in over 92% of cases when it is known *a priori* that just this species pair is expected.

Adjusting the prior probabilities of species membership according to a fly's wing tuft colouration does not improve allocation rate, so it is not recommended that this should be done.

Table 8.36

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|---------------------|---------------------|
| | <i>S. soubrense</i> | <i>S. squamosum</i> |
| 1 | 0.1905 | 0.8095 |
| 2 | 0.4062 | 0.5938 |
| 3 | 0.6655 | 0.3345 |
| 4 | 0.8526 | 0.1474 |
| 5 | 0.9439 | 0.0561 |

Table 8.37

Standardised Canonical Variate for *S. soubrense* and *S. squamosum*

| Character | Canonical Variate |
|---------------------|-------------------|
| Antennal Length 2 | -1.3938 |
| Antennal Segment 6 | -0.5038 |
| Wing Width 1 | 0.2542 |
| Wing Width 2 | 0.2544 |
| Tibia Length 1 | 0.9448 |
| Tibia Length 2 | -1.0839 |
| Basitarsus Length 2 | 1.1753 |

Table 8.38

Mean Vectors for species *S. soubrense* and *S. squamosum*

| Character | <i>S. soubrense</i> | <i>S. squamosum</i> |
|---------------------|---------------------|---------------------|
| Antennal Length 2 | 452.30756757 | 415.62558140 |
| Antennal Segment 6 | 45.50576577 | 40.36488372 |
| Wing Width 1 | 997.15797297 | 1048.84267442 |
| Wing Width 2 | 1401.13479279 | 1466.39627907 |
| Tibia Length 1 | 687.64756757 | 721.09465116 |
| Tibia Length 2 | 617.96972973 | 637.39744186 |
| Basitarsus Length 2 | 325.77297297 | 343.40406977 |

Table 8.39

Linear Discriminant functions for species *S. soubrense* and *S. squamosum*

| | <i>S. soubrense</i> | <i>S. squamosum</i> |
|----------|---------------------|---------------------|
| CONSTANT | -150.38773459 | -140.32654738 |
| V10 | 0.41885412 | 0.27668607 |
| V13 | 0.74360050 | 0.36064597 |
| V18 | 0.10132735 | 0.11313158 |
| V19 | 0.03829657 | 0.04664547 |
| V23 | 0.11749550 | 0.18100888 |
| V28 | -0.01654648 | -0.09953965 |
| V29 | -0.45363441 | -0.30004984 |

Table 8.40

Pooled within-species dispersion matrix for species *S. soubrense* and *S. squamosum*

| Character | V10 | V13 | V18 | V19 |
|-----------|---------------|---------------|---------------|---------------|
| V10 | 781.20558043 | 69.05704496 | 1353.66897306 | 1898.06680633 |
| V13 | 69.05704496 | 13.52189112 | 131.75112334 | 187.14544559 |
| V18 | 1353.66897306 | 131.75112334 | 4712.98607757 | 5870.89534929 |
| V19 | 1898.06680633 | 187.14544559 | 5870.89534929 | 9708.31406014 |
| V23 | 1022.42660720 | 94.10687451 | 2935.76237058 | 3942.77347389 |
| V28 | 933.35206049 | 87.13826351 | 2693.58657081 | 3612.26322503 |
| V29 | 530.81165688 | 52.03809484 | 1478.19712810 | 1991.04171904 |
| Character | V23 | V28 | V29 | |
| V10 | 1022.42660720 | 933.35206049 | 530.81165688 | |
| V13 | 94.10687451 | 87.13826351 | 52.03809484 | |
| V18 | 2935.76237058 | 2693.58657081 | 1478.19712810 | |
| V19 | 3942.77347389 | 3612.26322503 | 1991.04171904 | |
| V23 | 2287.88248971 | 1995.10209720 | 1090.85562028 | |
| V28 | 1995.10209720 | 1883.38427413 | 997.03216370 | |
| V29 | 1090.85562028 | 997.03216370 | 601.71878824 | |

8.3.1.9 Discrimination of *Simulium soubrense* and *S. yahense*

The 25 character set resulted in a Mahalanobis' squared distance of 5.25 (unbiased $D^2=4.82$, $e_{act}=0.136$) with 36 misallocated flies using resubstitution ($e_{res}=0.113$). A stepwise discriminant analysis resulted in an initial subset of 15 characters giving a Mahalanobis' squared distance of 5.02 (unbiased $D^2=4.77$, $e_{act}=0.137$) and 41 misallocated flies using resubstitution ($e_{res}=0.129$). Applying the dimension reduction method described in section 7.2.2 resulted in an eight character subset:

[V6, V10, V18, V22, V24, V25, V26, V29]

i.e. one head, one antennal, one wing and five leg measurement. This character subset resulted in a Mahalanobis' squared distance of 3.86 (unbiased $D^2=3.75$, $e_{act}=0.166$) and 41 misallocated flies ($e_{res}=0.129$), 31 *S. soubrense* classified as *S. yahense* and 10 *S. yahense* classified as *S. soubrense*.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in 33 misallocated flies (10.38%), 12 atypical flies (3.8%), 25 overlapping flies (7.86%) and 248 correctly allocated flies (77.99%).

The null hypothesis of equal wing tuft colouration and equal abdominal setal colouration were both rejected at $p<0.001$ using Wilcoxon two-sample rank sum tests. The prior probabilities of species membership according to a fly's wing tuft colouration were calculated and are shown as Table 8.41. When the prior probabilities were adjusted for wing tuft colouration the number of flies misallocated only fell to 34 ($e_{\text{res}}=0.1069$), with 29 *S. soubrense* classified as *S. yahense* and five *S. yahense* classified as *S. soubrense*. Including wing tuft colouration in the linear discriminant function resulted in 29 misallocated flies ($e_{\text{res}}=0.091$), 23 *S. soubrense* classified as *S. yahense*, and six *S. yahense* classified as *S. soubrense*. Including abdominal setal colouration either in the linear discriminant function or in the prior probabilities resulted in four flies being misallocated ($e_{\text{res}}=0.013$), these flies being the four *S. yahense* with abdominal setal colouration character state one. The special nature of this character is discussed in further detail in Chapter nine.

The null hypothesis of equal dispersion was not rejected at $p<0.001$ using the likelihood ratio test, therefore the dispersion matrices were pooled.

The standardised canonical variate (Table 8.42) shows that the main character involved in discrimination between these species is basitarsus length 2, with femur length 1 and tarsus segment 2 having a significant opposite influence.

The first eigenvector of the pooled within-species correlation matrix was a size vector with coefficients:

[0.35,0.32,0.35,0.37,0.37,0.36,0.34,0.37],

and accounted for 83.7% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 0.8193 to 0.6148 showing that size variation has some influence of the already relatively poor discrimination between these species.

The mean vectors are shown as Table 8.43, the linear discriminant functions as Table 8.44 and the pooled within-species dispersion matrix as Table 8.45.

To conclude, there is significant multivariate morphometric differentiation between *S. soubrense* and *S. yahense*, although this differentiation is not great. This differentiation involves both size and shape variation as in the absence of size the discrimination deteriorates. The eight character subset can be expected to classify flies to their correct species in nearly 80% of cases when it is known *a priori* that just this species pair is expected.

Adjusting the prior probabilities of species membership according to a fly's wing tuft colouration improves allocation rate, so it is recommended that this should be done in cases of doubt.

Table 8.41

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|---------------------|-------------------|
| | <i>S. soubrense</i> | <i>S. yahense</i> |
| 1 | 0.9832 | 0.0168 |
| 2 | 0.9343 | 0.0657 |
| 3 | 0.7756 | 0.2244 |
| 4 | 0.4564 | 0.5436 |
| 5 | 0.1694 | 0.8306 |

Table 8.42

Standardised Canonical Variate for *S. soubrense* and *S. yahense*

| Character | Canonical Variate |
|---------------------|-------------------|
| Head Width | -0.7411 |
| Antennal Length 2 | 0.2887 |
| Wing Width 1 | -0.2564 |
| Femur Length 1 | -1.2047 |
| Basitarsus Length 1 | 0.6478 |
| Tarsus Segment 2 | -1.1176 |
| Tarsus Segment 3 | 0.5920 |
| Basitarsus Length 2 | 2.5188 |

Table 8.43

Mean Vectors for species *S. soubrense* and *S. yahense*

| Character | <i>S. soubrense</i> | <i>S. yahense</i> |
|---------------------|---------------------|-------------------|
| Head Width | 806.88387027 | 828.68378125 |
| Antennal Length 2 | 452.30756757 | 475.44625000 |
| Wing Width 1 | 997.15797297 | 1039.91921875 |
| Femur Length 1 | 633.77135135 | 657.64000000 |
| Basitarsus Length 1 | 432.29513514 | 462.27500000 |
| Tarsus Segment 2 | 165.95333333 | 172.72166667 |
| Tarsus Segment 3 | 124.53063063 | 131.40125000 |
| Basitarsus Length 2 | 325.77297297 | 352.44218750 |

Table 8.44

Linear Discriminant functions for species *S. soubrense* and *S. yahense*

| | <i>S. soubrense</i> | <i>S. yahense</i> |
|----------|---------------------|-------------------|
| CONSTANT | -195.36976155 | -199.46621622 |
| V6 | 0.45053367 | 0.41909206 |
| V10 | 0.18662705 | 0.20430887 |
| V18 | 0.16447478 | 0.15714705 |
| V22 | -0.21177450 | -0.26513282 |
| V24 | -0.27046140 | -0.23166637 |
| V25 | 1.49677445 | 1.29398832 |
| V26 | -0.57435456 | -0.45030602 |
| V29 | -0.45106015 | -0.26045124 |

Table 8.45

Pooled within-species dispersion matrix for species *S. soubrense* and *S. yahense*

| Character | V6 | V10 | V18 | V22 |
|-----------|---------------|---------------|---------------|---------------|
| V6 | 2052.10239071 | 1005.73982336 | 2280.95932025 | 1675.78371658 |
| V10 | 1005.73982336 | 919.86051752 | 1327.76129138 | 986.45393206 |
| V18 | 2280.95932025 | 1327.76129138 | 4358.58590824 | 2476.85048452 |
| V22 | 1675.78371658 | 986.45393206 | 2476.85048452 | 1854.60560568 |
| V24 | 1150.43329966 | 694.13496194 | 1663.37658176 | 1158.69955525 |
| V25 | 373.05311287 | 225.45368038 | 550.82719596 | 383.35612405 |
| V26 | 306.66309748 | 172.83653605 | 462.67142999 | 305.37873041 |
| V29 | 879.53423597 | 515.24732941 | 1284.70297311 | 897.80567756 |
| VARIABLE | V24 | V25 | V26 | V29 |
| V6 | 1150.43329966 | 373.05311287 | 306.66309748 | 879.53423597 |
| V10 | 694.13496194 | 225.45368038 | 172.83653605 | 515.24732941 |
| V18 | 1663.37658176 | 550.82719596 | 462.67142999 | 1284.70297311 |
| V22 | 1158.69955525 | 383.35612405 | 305.37873041 | 897.80567756 |
| V24 | 889.95130489 | 272.87887342 | 212.26754456 | 642.00407298 |
| V25 | 272.87887342 | 107.98558312 | 79.66183143 | 212.12306725 |
| V26 | 212.26754456 | 79.66183143 | 78.24852773 | 164.49460988 |
| V29 | 642.00407298 | 212.12306725 | 164.49460988 | 525.87139075 |

8.3.1.10 Discrimination of *Simulium squamosum* and *S. yahense*

The Mahalanobis' squared distance between species using the 25 character set was 20.75 (unbiased $D^2=18.72$, $e_{act}=0.015$) with no mis-allocated flies using resubstitution ($e_{res}=0.0$). A stepwise

discriminant analysis resulted in an initial subset of 12 characters giving a Mahalanobis' squared distance between species of 20.88 (unbiased $D^2=19.86$, $e_{act}=0.013$) and 1 misallocated fly using resubstitution ($e_{res}=0.004$). Applying the method described in section 7.2.2 for dimension reduction in the context of discrimination, a six character subset resulted:

[V9,V11,V15,V22,V23,V29]

i.e. three antennal, and three wing characters. This character subset resulted in a Mahalanobis' squared distance of 16.6 (unbiased $D^2=16.2$, $e_{act}=0.022$) and three misallocated flies ($e_{res}=0.011$), 2 *S. squamosum* classified as *S. yahense* and 1 *S. yahense* classified as *S. squamosum*.

Allocation using typicality probability of species membership with atypicality defined at $\alpha=0.01$ resulted in three misallocated flies (1.12%), seven atypical flies (2.6%) three overlapping flies (1.12%) and 255 correctly allocated flies (95.15%).

The null hypothesis of equal wing tuft colouration and equal abdominal setal colouration were both rejected at $p<0.001$ using Wilcoxon two-sample rank sum tests. The prior probabilities of species membership according to a fly's wing tuft colouration were calculated and are shown as Table 8.46. When the prior probabilities were adjusted for wing tuft colouration 2 flies were misallocated using resubstitution ($e_{res}=0.075$), both of which were *S. squamosum* classified as *S. yahense*. Including wing tuft colouration in the linear discriminant function resulted in 1 misallocated fly ($e_{res}=0.0037$), a *S. squamosum* classified as *S. yahense*. Including abdominal setal colouration either in the linear discriminant function or in the prior probabilities resulted in four flies being misallo-

cated ($e_{res}=0.0149$), these flies being the four *S. yahense* with abdominal setal colouration character state one. The special nature of this character is discussed in further detail in Chapter nine.

The null hypothesis of equal dispersion was rejected at $p<0.001$ using the likelihood ratio test. However, for the practical and statistical reasons discussed in section 7.2.5, the dispersion matrices were still pooled.

The standardised canonical variate (Table 8.47) shows that antennal length 1 is the most important character involved in discrimination between these species, with tibia length 1 having a significant opposite influence.

The first eigenvector of the pooled within-species correlation matrix was a size vector with coefficients:

$$[0.40, 0.36, 0.38, 0.44, 0.44, 0.43],$$

and accounted for 78.1% of pooled within-species variation. When the scores along this vector were introduced into the model as a covariable the canonical root fell from 3.8454 to 3.6416 showing that size variation has negligible influence on discrimination between these species.

The mean vectors are shown as Table 8.48, the linear discriminant functions as Table 8.49 and the pooled within-species dispersion matrix as Table 8.50.

To conclude, there is significant multivariate morphometric differentiation between *S. squamosum* and *S. yahense*. This differentiation involves mainly shape variation because when size variation is controlled, discrimination is still effective. The six character subset can be expected to classify flies to their correct species in

over 95% of cases when it is known *a priori* that just this species pair is expected.

Adjusting the prior probabilities of species membership according to a fly's wing tuft colouration slightly improves allocation rate, so it is recommended that this should be done in cases of doubt following typicality probability allocation

Adjusting the prior probabilities of species membership according to a fly's abdominal setal colouration improves allocation rate but because of the problem that abdominal setal colouration is not 100% diagnostic for *S. yahense*, it is recommended that the character be used only in cases of doubt.

Table 8.46

Prior probabilities of species membership for each wing tuft category.

| Wing tuft category | Species | |
|--------------------|---------------------|-------------------|
| | <i>S. squamosum</i> | <i>S. yahense</i> |
| 1 | 1.0000 | 0.0000 |
| 2 | 0.9993 | 0.0007 |
| 3 | 0.7880 | 0.2120 |
| 4 | 0.0091 | 0.9909 |
| 5 | 0.0000 | 1.0000 |

Table 8.47

Standardised Canonical Variate for *S. squamosum* and *S. yahense*

| Character | Canonical Variate |
|--------------------|-------------------|
| Antennal Length 1 | 1.3978 |
| Antennal Segment 4 | 0.6915 |
| Antennal Segment 8 | 0.5655 |
| Femur Length 1 | -0.7261 |
| Tibia Length 1 | -0.9980 |
| Basitarsus Length | 0.4308 |

Table 8.48

Mean Vectors for species *S. squamosum* and *S. yahense*

| Character | <i>S. squamosum</i> | <i>S. yahense</i> |
|---------------------|---------------------|-------------------|
| Antennal Length 1 | 279.50755814 | 319.98281250 |
| Antennal Segment 4 | 37.69744186 | 45.00166667 |
| Antennal Segment 8 | 39.01674419 | 46.06083333 |
| Femur Length 1 | 654.56023256 | 657.64000000 |
| Tibia Length 1 | 721.09465116 | 721.60000000 |
| Basitarsus Length 2 | 343.40406977 | 352.44218750 |

Table 8.49

Linear Discriminant functions for species *S. squamosum* and *S. yahense*

| | <i>S. squamosum</i> | <i>S. yahense</i> |
|----------|---------------------|-------------------|
| CONSTANT | -134.68348826 | -171.69187356 |
| V9 | 0.80352471 | 1.02174383 |
| V11 | -0.47696294 | 0.06143848 |
| V15 | -0.20676689 | 0.28235314 |
| V22 | -0.10656273 | -0.16655907 |
| V23 | 0.12162377 | -0.04427324 |
| V29 | 0.15396694 | 0.22205447 |

Table 8.50

Pooled within-species dispersion matrix for species *S. squamosum* and *S. yahense*

| Character | V9 | V11 | V15 |
|-----------|---------------|---------------|---------------|
| V9 | 304.26573051 | 41.86146645 | 42.54871657 |
| V11 | 41.86146645 | 15.12832559 | 8.09739261 |
| V15 | 42.54871657 | 8.09739261 | 10.77909064 |
| V22 | 642.09971315 | 126.98207858 | 109.91556440 |
| V23 | 680.31886073 | 137.48406333 | 116.57001280 |
| V29 | 321.28657639 | 63.39587609 | 54.90745077 |
| Character | V22 | V23 | V29 |
| V9 | 642.09971315 | 680.31886073 | 321.28657639 |
| V11 | 126.98207858 | 137.48406333 | 63.39587609 |
| V15 | 109.91556440 | 116.57001280 | 54.90745077 |
| V22 | 2438.84470974 | 2485.28616321 | 1166.02930014 |
| V23 | 2485.28616321 | 2774.23815895 | 1244.06463814 |
| V29 | 1166.02930014 | 1244.06463814 | 648.06330373 |

8.3.2 OVERALL DISCRIMINATION

The full 25 character set excluding wing tuft colouration, abdominal setal colouration, and basitarsal spine number for the reasons given in section 7.2.3 resulted in the matrix of Mahalanobis' squared distances shown as Table 8.50a. All of these distances were significant at $p < 0.001$, but examining the matrix reveals that there are three morphometric subgroups within *S. damnosum* s.l.: *S. squamosum*, 'savanna', and *S. sanctipauli/S. soubrense/S. yahense*, with members of the last group being relatively close together.

The matrix of classifications using resubstitution is given as Table 8.51. The overall resubstituted error rate was 0.144. Of the individual species' error rates, *S. soubrense* is the highest with nearly 30% of flies being misallocated into other species.

A stepwise discriminant analysis on the 25 character set only rejected four characters, antennal segment 5, antennal segment 7, radial hair number and tarsus segment 3 (Chapter four). This was regarded as too large a character set for practical allocation purposes, and so the method for dimension reduction in the context of allocation described in Chapter seven was used. This resulted in a 13 character subset:

[V4, V9, V11, V13, V16, V17, V18, V19, V20, V22, V24, V28, V29]

i.e. one thorax, three antennal, five wing and four leg characters (Chapter four).

The matrix of Mahalanobis' squared distances between species resulting from these 13 characters is shown as Table 8.54. The table of resubstitutions (Table 8.55) shows that the resubstituted error rate for the 13 character subset was 0.15. The individual error rates range from 0.044 ('savanna') to 0.293 (*S. soubrense*).

Allocation using the typicality probability method given in section 7.2.4 resulted in 29 atypical flies (3.6%), 54 overlapping flies (6.75%), 94 incorrectly allocated flies (11.8%) and 623 correctly allocated flies (77.9%). Once the flies which lay on a species-pair overlap had been allocated using the appropriate species-pair statistics given in sections 8.3.1.1- 8.3.1.10, then 29 flies remained atypical (3.6%), 16 were still overlapping (2%), 104 were incorrect (13%) and 651 were correctly allocated (81.4%).

The null hypotheses of equal wing tuft colouration and equal abdominal setal colouration were both rejected at $p < 0.001$ using a Kruskal-Wallis test. Therefore the prior probabilities of species membership for each wing tuft colouration category and each abdominal setal colouration category were calculated using the method given in Chapter seven. These are shown as Tables 8.52 and 8.53.

When the prior probabilities were adjusted according to a fly's wing tuft colour, then the overall error rate remained at 0.15. When the priors were adjusted using both wing tuft colouration and abdominal setal colouration, then 89/799 flies were incorrectly allocated ($e_{res} = 0.108$).

The standardised canonical variates (Table 8.56) show the characters of importance in discrimination and also give each species' mean score along each of the canonical vectors as Table 8.57. The first canonical variate accounted for 76% of total variance and was influenced most by antennal length 1 and basitarsus length 2. The species best discriminated along this vector were *S. sanctipauli*/*S. yahense* at the positive end and 'savanna' at the negative end. The second canonical variate accounted for a further 17% of total variance, and was strongly positively influenced by basitarsus length 2

and wing length 3, and strongly negatively by thorax width and tibia length 2. Discrimination between *S. squamosum* at the positive end and *S. sanctipauli* at the negative end was the most important function of this vector. The third and fourth canonical vectors, whilst being statistically significant are probably unimportant biologically because both have small canonical roots (Campbell 1982).

The null hypothesis of equal dispersion was rejected at $p < 0.0001$ using the likelihood ratio test. However, for the practical and statistical reasons given in section 7.2.5 the dispersion matrices were still pooled.

The first principal component of the pooled within-species correlation matrix was a size vector with coefficients:

[0.29, 0.23, 0.19, 0.20, 0.30, 0.28, 0.29, 0.28, 0.30, 0.30, 0.30, 0.30, 0.30]

and accounted for 77.4% of pooled within-species variation. When the scores along this vector were included in the model as a covariable the canonical roots fell from:

[4.6377, 1.0643, 0.2908, 0.1459]

to

[4.052, 0.8732, 0.2502, 0.1366]

showing that size has a small influence on discrimination, principally along the second canonical variate.

The mean vectors are given as Table 8.58, the pooled dispersion matrix as Table 8.59 and the linear discriminant functions as Table 8.60.

To conclude, the 13 character subset derived in this analysis has shown that there is significant multivariate morphometric differentiation between the species of the *S. damnosum* complex. There is one major axis of between species variation, the first canonical variate,

which is a vector expressing antennal length and mid-leg basitarsus length in relation to the rest of the body. It is the relative size (i.e. shape) of these characters which is responsible for the discrimination between the species. Characters of secondary, but significant importance include thorax width, wing length, and the mid-leg tibia length. The other characters contribute to a lesser extent to between-species variation, probably through correlation with characters more important in species discrimination (Lubischew 1962).

Table 8.50a

Matrix of Mahalanobis' distances between species, 25 character set

| | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|-----------------------|---------|-----------------------|---------------------|---------------------|-------------------|
| Savanna | 0.0 | | | | |
| <i>S. sanctipauli</i> | 37.74 | 0.0 | | | |
| <i>S. soubrense</i> | 21.07 | 6.53 | 0.0 | | |
| <i>S. squamosum</i> | 11.91 | 25.89 | 12.91 | 0.0 | |
| <i>S. yahense</i> | 31.89 | 7.02 | 6.01 | 14.6 | 0.0 |

Table 8.51

Table of re-classifications, using resubstitution, 25 character set

| | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|-----------------------|---------|-----------------------|---------------------|---------------------|-------------------|
| Savanna | 241 | 0 | 3 | 5 | 0 |
| <i>S. sanctipauli</i> | 0 | 51 | 5 | 0 | 5 |
| <i>S. soubrense</i> | 7 | 21 | 156 | 7 | 31 |
| <i>S. squamosum</i> | 9 | 0 | 4 | 158 | 1 |
| <i>S. yahense</i> | 0 | 10 | 7 | 0 | 79 |

$$e_{res} = 115/800 = 0.1438$$

Table 8.52

Prior probability of species membership for each wing tuft colour category

| Wing Tuft Colour | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|------------------|---------|-----------------------|---------------------|---------------------|-------------------|
| 1 | 0.5644 | 0.0009 | 0.0320 | 0.4026 | 0.0000 |
| 2 | 0.2766 | 0.0273 | 0.2357 | 0.4580 | 0.0024 |
| 3 | 0.0395 | 0.2432 | 0.5063 | 0.1519 | 0.0591 |
| 4 | 0.0012 | 0.4534 | 0.2273 | 0.0105 | 0.3076 |
| 5 | 0.0000 | 0.3316 | 0.0400 | 0.0003 | 0.6280 |

Table 8.53

Prior probability of species membership for each wing tuft and abdominal setal colour category

| A ¹ | B ² | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|----------------|----------------|---------|-----------------------|---------------------|---------------------|-------------------|
| 1 | 1 | 0.5645 | 0.001 | 0.0319 | 0.4026 | 0.0 |
| 1 | 2 | 0.2768 | 0.0285 | 0.2359 | 0.4587 | 0.0 |
| 1 | 3 | 0.0417 | 0.263 | 0.535 | 0.1604 | 0.0 |
| 1 | 4 | 0.0017 | 0.6551 | 0.3281 | 0.0152 | 0.0 |
| 1 | 5 | 0.0 | 0.8895 | 0.1097 | 0.0008 | 0.0 |
| 2 | 1-5 | 0.0 | 0.0 | 0.0 | 0.0 | 1.0 |

A¹=Abdominal setal colour category (see Chapter four).

B²=Wing tuft colour category (see Chapter four).

Table 8.54

Matrix of Mahalanobis' distances between species, 13 character subset

| | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|-----------------------|---------|-----------------------|---------------------|---------------------|-------------------|
| Savanna | 0.0 | | | | |
| <i>S. sanctipauli</i> | 35.85 | 0.0 | | | |
| <i>S. soubrense</i> | 19.85 | 5.89 | 0.0 | | |
| <i>S. squamosum</i> | 10.79 | 23.48 | 11.04 | 0.0 | |
| <i>S. yahense</i> | 29.47 | 6.08 | 4.91 | 12.58 | 0.0 |

Table 8.55

Table of re-classifications, using resubstitution, 13 character subset

| | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|-----------------------|---------|-----------------------|---------------------|---------------------|-------------------|
| Savanna | 238 | 0 | 2 | 9 | 0 |
| <i>S. sanctipauli</i> | 0 | 48 | 5 | 0 | 8 |
| <i>S. soubrense</i> | 6 | 19 | 157 | 10 | 30 |
| <i>S. squamosum</i> | 7 | 0 | 4 | 159 | 2 |
| <i>S. yahense</i> | 0 | 10 | 8 | 0 | 78 |

$$e_{res} = 120/800 = 0.15$$

Table 8.56

Standardised Canonical Variates

| Character | CV I | CV II | CV III | CV IV |
|--------------------|---------|---------|---------|---------|
| Thorax Width | -0.6323 | -1.1533 | -0.0874 | -0.7891 |
| Antennal Length 1 | 1.6678 | -0.4761 | -0.1345 | -0.5461 |
| Antennal Segment 4 | 0.1762 | -0.3361 | 0.8469 | 0.1080 |
| Antennal Segment 6 | 0.6211 | -0.3631 | -0.6173 | 0.3769 |
| Wing Length 1 | -0.1310 | -0.6600 | 0.4007 | 1.0422 |
| Wing Length 2 | -0.3995 | 0.1800 | 0.0109 | -0.8774 |
| Wing Width 1 | -0.1248 | 0.3747 | -0.5498 | 1.0485 |
| Wing Width 2 | -0.2206 | -0.1079 | 1.2190 | 0.1655 |
| Wing Length 3 | 0.2579 | 1.1661 | -0.4448 | -2.4018 |
| Femur Length 1 | 0.0959 | 0.8137 | -2.5897 | 1.0942 |
| Basitarsus Length | -0.5418 | 0.1549 | 0.8505 | 1.1273 |
| Tibia Length 2 | -0.0902 | -1.6104 | 0.8879 | -0.0030 |
| Basitarsus Length | 1.1033 | 2.1201 | 0.6605 | -0.1962 |

Table 8.57

Species Means on Canonical Variates

| Species | CV I | CV II | CV III | CV IV |
|-----------------------|---------|---------|---------|---------|
| Savanna | -2.7399 | -0.6702 | 0.1969 | -0.0604 |
| <i>S. sanctipauli</i> | 3.1093 | -1.0771 | 0.8367 | 0.9707 |
| <i>S. soubrense</i> | 1.6280 | -0.6058 | -0.6789 | -0.1120 |
| <i>S. squamosum</i> | -0.6053 | 1.7760 | -0.1843 | 0.2615 |
| <i>S. yahense</i> | 2.4507 | 0.6415 | 0.8578 | -0.6699 |

Table 8.58

Mean Vectors

| Character | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> |
|---------------------|---------------|-----------------------|---------------------|
| Thorax Width | 867.74462410 | 909.84725902 | 872.95284865 |
| Antennal Length 1 | 254.11867470 | 328.20737705 | 304.93783784 |
| Antennal Segment 4 | 35.31759036 | 48.33967213 | 42.19909910 |
| Antennal Segment 6 | 36.08449799 | 50.53508197 | 45.50576577 |
| Wing Length 1 | 708.12433735 | 774.45639344 | 730.55351351 |
| Wing Length 2 | 439.08530120 | 461.91540984 | 443.02162162 |
| Wing Width 1 | 978.44353414 | 1041.64434426 | 997.15797297 |
| Wing Width 2 | 1389.53391566 | 1479.41942623 | 1401.13479279 |
| Wing Length 3 | 1414.92809237 | 1522.32909836 | 1478.64736486 |
| Femur Length 1 | 602.69012048 | 658.95737705 | 633.77135135 |
| Basitarsus Length 1 | 419.64240964 | 461.10885246 | 432.29513514 |
| Tibia Length 2 | 601.62313253 | 650.97245902 | 617.96972973 |
| Basitarsus Length 2 | 307.67710843 | 347.06803279 | 325.77297297 |

Table 8.58 (continued)

Mean Vectors

| Character | <i>S. squamosum</i> | <i>S. yahense</i> |
|---------------------|---------------------|-------------------|
| Thorax Width | 891.07350698 | 906.62228125 |
| Antennal Length 1 | 279.50755814 | 319.98281250 |
| Antennal Segment 4 | 37.69744186 | 45.00166667 |
| Antennal Segment 6 | 40.36488372 | 47.19750000 |
| Wing Length 1 | 758.70976744 | 770.69750000 |
| Wing Length 2 | 463.22372093 | 471.03875000 |
| Wing Width 1 | 1048.84267442 | 1039.91921875 |
| Wing Width 2 | 1466.39627907 | 1576.95807292 |
| Wing Length 3 | 1540.09226744 | 1495.92739583 |
| Femur Length 1 | 654.56023256 | 657.64000000 |
| Basitarsus Length 1 | 453.87000000 | 462.27500000 |
| Tibia Length 2 | 637.39744186 | 648.41500000 |
| Basitarsus Length 2 | 343.40406977 | 352.44218750 |

Table 8.59

Pooled within-species dispersion matrix

| Character | V4 | V9 | V11 | V13 |
|-----------|---------------|--------------|--------------|--------------|
| V4 | 4020.42684612 | 674.11534770 | 123.93425781 | 114.81760517 |
| V9 | 674.11534770 | 290.36670606 | 36.22041760 | 37.77400930 |
| V11 | 123.93425781 | 36.22041760 | 14.12867695 | 9.20617799 |
| V13 | 114.81760517 | 37.77400930 | 9.20617799 | 12.62327975 |
| V16 | 2819.63025059 | 573.81903537 | 94.33719941 | 96.15314058 |
| V17 | 1705.33879012 | 349.30148288 | 55.26046145 | 56.36812476 |
| V18 | 3479.02054715 | 645.92634072 | 109.26739900 | 113.81280202 |
| V19 | 4779.96597811 | 947.15089526 | 156.22669013 | 162.58573156 |
| V20 | 4901.55169491 | 998.76130192 | 163.94461750 | 168.26927707 |
| V22 | 2430.36654352 | 479.44726851 | 81.94251823 | 81.64578963 |
| V24 | 1660.98659993 | 332.33085966 | 58.78564509 | 57.16285608 |
| V28 | 2354.07338731 | 450.54402160 | 76.77856170 | 76.01723060 |
| V29 | 1262.94455076 | 250.45044613 | 43.66888838 | 43.47540170 |

| Character | V16 | V17 | V18 | V19 |
|-----------|---------------|---------------|---------------|---------------|
| V4 | 2819.63025059 | 1705.33879012 | 3479.02054715 | 4779.96597811 |
| V9 | 573.81903537 | 349.30148288 | 645.92634072 | 947.15089526 |
| V11 | 94.33719941 | 55.26046145 | 109.26739900 | 156.22669013 |
| V13 | 96.15314058 | 56.36812476 | 113.81280202 | 162.58573156 |
| V16 | 2754.94213103 | 1510.11756107 | 2949.81478407 | 3918.69270629 |
| V17 | 1510.11756107 | 1048.71143120 | 1678.25278978 | 2414.76598315 |
| V18 | 2949.81478407 | 1678.25278978 | 4112.98487223 | 5067.96192179 |
| V19 | 3918.69270629 | 2414.76598315 | 5067.96192179 | 8195.58351299 |
| V20 | 4243.34067646 | 2465.36734418 | 5134.02817590 | 7155.46508249 |
| V22 | 1999.44104911 | 1188.65965191 | 2423.13912677 | 3276.69614499 |
| V24 | 1380.11258810 | 836.14006698 | 1641.46026172 | 2268.03800475 |
| V28 | 1953.41437341 | 1168.40568665 | 2362.22639952 | 3208.22684180 |
| V29 | 1061.19383404 | 629.00740560 | 1268.70992201 | 1727.30826743 |

| Character | V20 | V22 | V24 | V28 |
|-----------|---------------|---------------|---------------|---------------|
| V4 | 4901.55169491 | 2430.36654352 | 1660.98659993 | 2354.07338731 |
| V9 | 998.76130192 | 479.44726851 | 332.33085966 | 450.54402160 |
| V11 | 163.94461750 | 81.94251823 | 58.78564509 | 76.77856170 |
| V13 | 168.26927707 | 81.64578963 | 57.16285608 | 76.01723060 |
| V16 | 4243.34067646 | 1999.44104911 | 1380.11258810 | 1953.41437341 |
| V17 | 2465.36734418 | 1188.65965191 | 836.14006698 | 1168.40568665 |
| V18 | 5134.02817590 | 2423.13912677 | 1641.46026172 | 2362.22639952 |
| V19 | 7155.46508249 | 3276.69614499 | 2268.03800475 | 3208.22684180 |
| V20 | 8224.10690353 | 3472.63849913 | 2387.76530057 | 3377.77197174 |
| V22 | 3472.63849913 | 1816.16396259 | 1154.99263296 | 1670.50349394 |
| V24 | 2387.76530057 | 1154.99263296 | 881.10960971 | 1134.99996767 |
| V28 | 3377.77197174 | 1670.50349394 | 1134.99996767 | 1706.41061326 |
| V29 | 1835.20200128 | 887.63622455 | 635.37505416 | 870.84426109 |

Table 8.59(continued)

| Character | V29 |
|-----------|---------------|
| V4 | 1262.94455076 |
| V9 | 250.45044613 |
| V11 | 43.66888838 |
| V13 | 43.47540170 |
| V16 | 1061.19383404 |
| V17 | 629.00740560 |
| V18 | 1268.70992201 |
| V19 | 1727.30826743 |
| V20 | 1835.20200128 |
| V22 | 887.63622455 |
| V24 | 635.37505416 |
| V28 | 870.84426109 |
| V29 | 516.02553624 |

Table 8.60

Linear Discriminant functions

| | Savanna | <i>S. sanctipauli</i> | <i>S. soubrense</i> | <i>S. squamosum</i> | <i>S. yahense</i> |
|-------|-------------|-----------------------|---------------------|---------------------|-------------------|
| CONS- | -159.733167 | -221.51972253 | -196.68139491 | -187.19981965 | -219.46907820 |
| -TANT | | | | | |
| V4 | -0.02705877 | -0.09002968 | -0.06882842 | -0.09453154 | -0.09422150 |
| V9 | 0.55777538 | 0.85369529 | 0.79324788 | 0.63004027 | 0.82112976 |
| V11 | 0.15679247 | 0.48019491 | 0.15689796 | 0.02635963 | 0.32868239 |
| V13 | -0.21406578 | 0.41275768 | 0.31915059 | -0.08214442 | 0.13634649 |
| V16 | -0.14728705 | -0.13290588 | -0.16488123 | -0.17682596 | -0.18035777 |
| V17 | 0.04504152 | -0.05064176 | -0.00406664 | 0.02482137 | 0.00756623 |
| V18 | 0.15185719 | 0.14968480 | 0.15052823 | 0.16886344 | 0.13539866 |
| V19 | 0.04640510 | 0.04342813 | 0.02582277 | 0.03487736 | 0.04055290 |
| V20 | 0.10955495 | 0.09344627 | 0.12554969 | 0.13566331 | 0.14721648 |
| V22 | -0.13414412 | -0.14034412 | -0.07841564 | -0.06072305 | -0.15103736 |
| V24 | 0.05064384 | 0.00576679 | -0.04229086 | 0.02890527 | -0.02977846 |
| V28 | 0.14400142 | 0.15937790 | 0.11569554 | 0.04474148 | 0.09978495 |
| V29 | -0.34843142 | -0.14215441 | -0.19269392 | -0.09191129 | -0.02668580 |

8.4 DISCUSSION

The 'global' discriminant analyses presented in this chapter have shown that there is a large amount of multivariate morphometric variation between females of the five taxa examined within the *S. damnosum* complex from West Africa. These analyses also demonstrate that multivariate statistical methods contribute valuable information to the morphological study of the complex, and should provide an additional, powerful method for the field identification of flies.

The species which is most successfully discriminated is 'savanna', the pooled *S. damnosum s.s.*, *S. sirbanum* category. This is true of both the species-pair analyses and the overall analysis. The species with which it has most phenetic overlap is *S. squamosum*, but this is only about 5%. The ability to identify this species with accuracy is of considerable importance, because both *S. damnosum s.s.* and *S. sirbanum* are known to be dangerous vectors of the more debilitating savanna strain of *O. volvulus*. The rate of correct identification using the multivariate statistical method is an improvement over published methods currently in routine use for adult female identification.

The morphological characters which were most commonly derived in discriminant analyses involving 'savanna' were thorax width, antennal length and mid-basitarsus length (Chapter four) for all of which it is relatively smaller than the other species. 'Savanna' flies are also pale, mostly having wing tuft colour category one and two, and always showing abdominal setal colour category two.

Size variation was a significant component of the between-species analyses involving 'savanna', but the size-free canonical roots were still higher than for most other analyses, demonstrating that both

size and shape differences are important in defining the distinctiveness of this taxon. Whether the consistently smaller size of 'savanna' flies is a taxonomic feature, or whether it is an environmentally mediated feature which might disappear once a wider range of geographic and seasonal variation has been sampled cannot be stated on the available data.

Simulium squamosum is the next most isolated species, with phenetic overlap ranging from about 3% to 8%. The species to which it is phenetically closest is 'savanna', followed by *S. soubrense* and *S. yahense*, and it is morphologically most distant from *S. sanctipauli*. It is of interest that *S. yahense*, to which it is chromosomally very close (Vajime and Dunbar 1975) is morphologically so distant. Clearly, morphological and chromosomal evolution in the *S. damnosum* complex are not always closely correlated, as was emphasised in Chapter six.

Morphological characters which are taxonomically useful in distinguishing *S. sanctipauli* in relation to the other species include having a relatively broader thorax, and longer wing, and the leg characters. The species is also paler than most, except for 'savanna'.

The ability to identify this species using multivariate morphometrics is an improvement over the currently used morphological methods (e.g. Garms and Cheke 1985), but the rate of correct identification is not as good as that achieved using enzyme electrophoresis (Meredith and Townson 1981). In combination, the two methods should be able to identify flies with great accuracy, assuming the rate of incorrect identification for each method is independent.

The three species *S. soubrense*, *S. sanctipauli* and *S. yahense* show a larger degree of phenetic overlap. The morphological similarity of *S. soubrense* and *S. sanctipauli* might have been expected considering their chromosomal relatedness, but the proximity of *S. yahense* to this pair of species is unusual considering their chromosomal distance. Wing tuft colour parallels this morphological similarity, with *S. yahense* and *S. sanctipauli* being the darkest and *S. soubrense* being very variable (Chapter six). Abdominal setal colour is 96% diagnostic for *S. yahense* on the basis of this data. The cause of the phenetic similarity of the three species may be due to ecological similarities between them, as all three are predominantly forest dwellers (although *S. soubrense* is variable).

Table 8.61 summaries the results of the overall discrimination using the five allocation schemes used in this analysis, forced allocation without adjusted prior probabilities, forced allocation with priors adjusted for wing tuft colour, forced allocation with priors adjusted for wing tuft and abdominal setal colour, and typicality probability allocation with and without subsequent allocation of overlapping flies using the appropriate species-pair statistics.

The method which results in the largest number of correct allocations is forced allocation with priors adjusted according to a fly's wing tuft and abdominal setal colour, although comparison of the first two columns reveals that this is due entirely to abdominal setal colour. By automatically classifying all category 2 flies into *S. yahense* and all category 1 flies into some other species, the character is acting as a diagnostic character in traditional taxonomic analysis.

Typicality probability allocation without subsequent allocation of overlapping flies results in considerably fewer correct allocations, although as was emphasised in Chapter seven, the purpose of using this method is to provide a biologically more meaningful method of allocation. Ignoring atypicality or overlapping flies will ultimately lead ^{to} poorer allocation rates.

Table 8.61

Comparison of five methods of allocation

| | Forced ¹ | Forced ² | Forced ³ | Typicality | Typicality |
|-------------|---------------------|---------------------|---------------------|------------|------------|
| Correct | 680 | 680 | 713 | 623 | 651 |
| Incorrect | 120 | 120 | 86 | 94 | 104 |
| Overlapping | na | na | na | 54 | 16 |
| Atypical | na | na | na | 29 | 29 |

¹Forced allocation without adjusted priors.

²Forced allocation with prior probabilities adjusted for wing tuft colour

³Forced allocation with prior probabilities adjusted for wing tuft colour and abdominal setal colour

⁴Typicality probability without subsequent species pair allocation of overlapping flies

⁵Typicality probability with subsequent species pair allocation of overlapping flies

CHAPTER NINE: GENERAL DISCUSSION

9.1 INTRODUCTION

The previous chapters of this project have examined variation within and between species of the *S. damnosum* complex in West Africa from three distinct viewpoints:

1. 'Classical' larval polytene chromosome analysis,
2. Multivariate statistical analysis of larval polytene chromosome variation,
3. Multivariate statistical analysis of adult female morphological variation.

The purpose of this chapter is to discuss issues arising from these three approaches to the analysis of variation, to summarise and discuss the adult morphological variation, to provide worked examples of the mathematics of the different allocation procedures for identifying females, and to suggest a protocol for the field identification of adult females. Suggestions for future study will also be made.

9.2 CHROMOSOMAL VARIATION

The study of larval polytene chromosome variation in the *S. damnosum* complex using the 'classical' methods (e.g. Vajime and Dunbar 1975, Quillevére 1975, Post 1986) is justified because species which are responsible for more serious disease transmission (e.g. *S. sirbanum*, *S. damnosum s.s.*) can be identified, and also species can be associated with ecological and bionomic factors which can influence control of the vector. Vector control problems such as the evolution of insecticide resistance and reinvasion of flies from uncontrolled areas can be clarified by detailed cytotaxonomic study of the problem flies, which in turn helps to overcome such problems more quickly than if the species complex were regarded as unitary.

The identification of a new chromosomal form, *S. sanctipauli* 'Djodji' from Togo (Chapter two, Surtees *et al.* 1988) demonstrates the continuing value of the usual approach to polytene chromosome variation, and the epidemiological work arising from its recognition (Cheke and Denke 1988) justifies the detailed analysis of chromosomal variation within previously recognised species.

However, the multivariate analyses presented in Chapter three demonstrate clearly that chromosomal variation in the *S. damnosum* complex is much more subtle and involved than can be revealed simply by using the classical approach. Important substantive results arising from this analysis include the recognition that the insecticide resistant flies within the *S. sanctipauli* subcomplex are chromosomally distinctive in relation to other *S. sanctipauli*, and may represent a new chromosomal form, a finding which had been overlooked using classical cytotaxonomy. That the variation revealed within *S. soubrense* 'Menankaya/Konkoure' cannot simply be broken into

classically defined taxa without considerable loss of information is another substantive result of this analysis, as is the recognition of the chromosomal distinctiveness of *S. soubrense* 'B'. Apart from these substantive findings, however, the importance of the analyses shown in Chapter three lies in the new approach taken to the analysis of polytene chromosome variation within the *S. damnosum* complex. As emphasised in that chapter, the methods used and the materials available were not entirely optimal, nevertheless, it is clear that the method applied routinely to the *S. damnosum* complex will reveal greater understanding of the population structure of the different species which will include factors of importance in vector control and disease transmission such as tracing insecticide resistance, more detailed and subtle analysis of relative rates of gene flow and migration, and the recognition of new sibling species.

For the approach adopted in Chapter three to be applied to other parts of the *S. damnosum* complex, and to the complex as a whole, the homologies of chromosomal sequences amongst different taxa must first be established. It is known that this work is not complete (Post personal communication), and until it is, the method will be restricted to better known parts of the complex such as the *S. sanctipauli* subcomplex and the *S. squamosum* subcomplex.

9.3 ADULT FEMALE MORPHOLOGICAL VARIATION

The adult female morphological characters described and figured in Chapter four were used to examine intra- and inter-specific variation in the *S. damnosum* complex, with the results of these analyses presented in Chapters six, seven and eight.

The purpose of this section is to examine which of the morphological characters were most important in discrimination, to discuss the significance of size variation in the *S. damnosum* complex, to comment on the two colour characters and to discuss the importance of the departure from the null assumption of equal dispersion.

9.3.1 MORPHOLOGICAL CHARACTERS OF IMPORTANCE

Table 9.1 summarises the characters derived for each of the analyses presented in Chapters six, seven and eight. The first column refers to the chapter heading for each analysis.

9.3.1.1 Thorax Characters

Of the two thorax characters measured in this project, thorax width (V4) was more important taxonomically than thorax length (V3). Neither character was important in describing intraspecific variation (analyses 6.3.1-6.3.6), but of the 35 between-species analyses, thorax length was derived in six analyses, and thorax width in 16. Of these, none of the standardised canonical variate coefficients were relatively large for thorax length whereas 10 were considered important for thorax width. Thus thorax width seems to be a useful taxonomic character in the *S. damnosum* complex.

9.3.1.2 Head Characters

Of the two head characters used in this project, vertex width (V5) was not derived in any of the discriminant analyses, and so can be

discounted as a taxonomic character. Head width (V6) was selected in 12 of the analyses, but was only considered important in the within-species analysis for *S. sirbanum*, and in the species pair discrimination of *S. soubrense* 'Beffa' and *S. sanctipauli* 'Djodji'. Therefore, neither character on its own can be considered of great importance taxonomically, although head width can be useful in conjunction with other characters.

9.3.1.3 Antennal Characters

Seven antennal characters were measured in this project, and they contributed considerably to the discrimination between species.

Antennal length one (V9), which is the shorter of the two measurements which included more than one segment, was selected in 21 of the analyses, while antenna length two (V10) was selected in 12 of the analyses. In no analysis were both selected together. Of the analyses in which they were selected, antennal length one was considered important in 14 and antennal length two in nine.

Of the within-species analyses, only in *S. sanctipauli* was antennal length considered important, indicating that the character may be relatively invariant within species, but of considerable importance between species.

The individual antennal segments 4-8 (V11-V15) were sometimes derived in the between-species analyses, but the only segment considered important was antennal segment 6 (V13), which was chosen in analyses including 'savanna' species and members of the *S. sanctipauli* subcomplex.

Antennal length and the relative compaction of the antennal segments 4-8 were considered important in previous morphological studies

of the *S. damnosum* complex (e.g. Garms 1978, Quillévére *et al.* 1977) and this has been confirmed in these analyses.

9.3.1.4 Wing Characters

Excluding wing tuft colour, which is dealt with in section 9.3.3, six characters were measured on the wing. Five of these were linear measurements, and one a count. This character, the number of hairs on the radial vein of the wing (V21) was chosen initially because it had previously been considered taxonomically important by Quillévére and Sechan (1978). However it was derived in the characters subsets of only one of the 35 analyses in chapters six, seven and eight, and it was not considered important in this analysis. Therefore, the character can be discounted as a taxonomic character in the *S. damnosum* complex, confirming the findings of Garms (1978) and Townson and Meredith (1979).

Of the linear measurements, the longest wing length measurement (V20) was chosen in 24 of the 35 analyses and considered particularly important on three of the six within-species analyses and 14 of the 29 between-species analyses. Thus wing length appears to be a taxonomically useful character.

Of the other characters, V17 and V18 were quite frequently present in the final character subsets of the between-species analyses.

9.3.1.5 Leg Characters

Eight leg character were measured (excluding basitarsus length 2 which was rejected at an early stage in the analysis), five on the fore-leg, three on the mid-leg. The least important characters on the fore-leg were the very small tarsal segments. This is possibly a reflection of their small size decreasing the accuracy of measurement.

Femur length, tibia length and basitarsal length of the fore-leg were sometimes considered important in between-species analyses, but none were of great taxonomic value.

Of the three mid-leg characters, the least important was femur length, but tibia length and basitarsal length were frequently considered important, and were often chosen together, indicating that the relative proportion of these segments might be of taxonomic significance in the *S. damnosum* complex.

9.3.2 INFLUENCE OF SIZE VARIATION

Table 9.2 summarises the influence of size variation on the within- and between-species analyses presented in Chapters six, seven and eight.

The proportion of variance along the first principal component of the pooled within-groups correlation indicates the amount of error variance due to size. This proportion ranges from 58.6% to 88.9% showing that within-groups variation is predominantly along the size axis, i.e. individuals within a sample, or within a species are more likely to differ in size than shape.

The size-free canonical root is an indication of the shape-only differentiation between samples (6.3.1.1-6.3.1.6) or between species (7.3.1.1-8.3.2). As an informal guideline based on extensive practical experience, Campbell (1978) suggests that a canonical root less than about 0.75 to 1.0 is unlikely to be of practical use, although this guideline was not for size-free canonical roots. Using this guideline strictly and conservatively, only one of the within-species analyses (*S. soubrense*) shows meaningful shape differentiation, reflecting the chromosomal heterogeneity of this taxon.

Six of the between-species analyses have size-free canonical roots less than one, these are *S. damnosum s.s./S. sirbanum*, *S. soubrense 'B'/S. soubrense*, *S. soubrense 'B'/S. yahense*, *S. soubrense/S. yahense* and *S. sanctipauli/S. soubrense*. All other analyses have shape-only canonical roots greater than one and so in principle all are useful.

The influence of size variation, as measured by the ratio of the size-free canonical root to the size-in canonical root, expressed as a percentage, gives the influence of the pooled within-groups scatter along the size vector on the canonical variate analysis. This influence varies greatly, from very heavy influence (33.2%) to negligible influence (97.9%).

There appears to be no obvious pattern as to the influence of within-groups size variation on between-groups discrimination, beyond seeing that 'savanna' flies tend to be smaller. It seems that size variation is a random component of within-species variation which does not influence the taxonomically more important shape differences between species.

To conclude, whilst size variation is the predominant mode of variation in *S. damnosum s.l.*, the extent of between-species shape differentiation is unaffected by size variation.

Table 9.1

Characters of Importance

| Section | Thorax | Head | Antenna | Wing | Fore-leg | Mid-leg |
|----------|---------------|--------------|-----------------------|--------------------|--------------|-----------------------|
| 6.3.1 | | | V9 | <u>V16</u> | V22 V25 | V27 |
| 6.3.2 | V3 V4 | V6 | | V17 <u>V18</u> V20 | <u>V23</u> | |
| 6.3.3 | | | <u>V9</u> | V20 | | |
| 6.3.4 | V4 | | V10 | V17 <u>V18</u> V20 | V24 | V27-V29 |
| 6.3.5 | | | V9 V11 | V19 V20 | V22 | V28 <u>V29</u> |
| 6.3.6 | | V6 | | V17V19V20 | V22 | <u>V27</u> <u>V28</u> |
| 7.3.1.1 | | | <u>V9</u> V14 | V18 V19 | | V27 <u>V28</u> |
| 7.3.1.2 | | <u>V6</u> | <u>V10</u> | V17 <u>V20</u> | | V27 <u>V29</u> |
| 7.3.1.3 | V6 <u>V10</u> | <u>V18</u> V | 0 | | | |
| 7.3.1.4 | | V6 | <u>V9</u> <u>V13</u> | | | V29 |
| 7.3.1.5 | V3 <u>V4</u> | | V9 V12 | V17 V19 | <u>V23</u> | <u>V28</u> <u>V29</u> |
| 7.3.1.6 | | | V10 | V18 | V22 | |
| 7.3.2 | <u>V4</u> | V6 | <u>V9</u> | <u>V18</u> V20 | <u>V23</u> | V27V28V29 |
| 7.4.1.1 | V4 | | V9 V14 | V16 <u>V17</u> | V23 | <u>V29</u> |
| 7.4.1.2 | | | <u>V9</u> | V17 | | V29 |
| 7.4.1.3 | V3 | | <u>V9</u> V13 | <u>V19</u> | | V29 |
| 7.4.1.4 | <u>V4</u> | | <u>V10</u> <u>V13</u> | V19 | | V29 |
| 7.4.1.5 | <u>V4</u> | V6 | V9 | V20 | | V27V28V29 |
| 7.4.1.6 | <u>V4</u> | | V9 | V20 | | V29 |
| 7.4.1.7 | <u>V4</u> | | V10 V15 | V16 V20 | V24 | V29 |
| 7.4.1.8 | | | V11 V15 | V19 <u>V20</u> | | <u>V27</u> |
| 7.4.1.9 | | | <u>V11</u> | V20 | | |
| 7.4.1.10 | V3 | | V9 V11 | V17 V20 | V22 | |
| 7.4.1.11 | V3 V4 | | | V16V17V20V21 | V22 | <u>V28</u> |
| 7.4.1.12 | V4 | | V10 | V17 V20 | | |
| 7.4.1.13 | <u>V4</u> | | V10 | V17V19V20 | | V27V28V29 |
| 7.4.1.14 | V3 V4 | V6 | <u>V10</u> V13 | | V24 | V28 <u>V29</u> |
| 7.4.1.15 | | V6 | | V17 V18 | V22V24V25V26 | V29 |
| 7.4.1.16 | | | V9V11V15 | V16 | V22 V24 | <u>V29</u> |
| 7.4.2 | V4 | V6 | <u>V9</u> V11 | V16V17V19V20 | V22 | V28 <u>V29</u> |
| 8.3.1.1 | <u>V4</u> | | <u>V9</u> <u>V13</u> | | | V29 |
| 8.3.1.2 | | | <u>V9</u> V13 | V17 V19 | V22 | |
| 8.3.1.3 | V4 | | V9 V12 | V20 | | V28 <u>V29</u> |
| 8.3.1.4 | <u>V4</u> | | <u>V9</u> | V20 | | V29 |
| 8.3.1.5 | | V6 | <u>V14</u> V15 | V16V17V19V20 | V24 | V27 |
| 8.3.1.6 | | | V10 V13 | V16 V18 V20 | V22 | |
| 8.3.1.7 | V3 | | V11 | V16 <u>V17</u> V20 | | <u>V28</u> |
| 8.3.1.8 | | | <u>V10</u> V13 | V18 V19 | V23 | V28 <u>V29</u> |
| 8.3.1.9 | | V6 | V10 | V18 | V22V24V25V26 | V29 |
| 8.3.1.10 | | | V9 V11 V15 | | V22 V23 | V29 |
| 8.3.2 | <u>V4</u> | | <u>V9</u> V11 V13 | V16-V20 | V22 V24 | V28 <u>V29</u> |

Table 9.2
The influence of size variation

| Analysis ¹ | % Var Size ² | Size-free Root ³ | Size Influence ⁴ | Species ⁵ |
|-----------------------|-------------------------|-----------------------------|-----------------------------|----------------------|
| 6.3.1.1 | 74 | 0.717, 0.122 | 79.6, 53.4 | dam |
| 6.3.1.2 | 85.0 | 0.811, 0.243 | 69.5, 95.8 | sir |
| 6.3.1.3 | 69.5 | 0.828 | 99.84 | san |
| 6.3.1.4 | 83.7 | 1.829, 1.206, | 84.3, 76 | soub |
| 6.3.1.5 | 69.6 | 0.927, 0.348 | 39.5, 35.3 | squ |
| 6.3.1.6 | 86 | 0.453, 0.364 | 35.3, 35.3 | yah |
| 7.3.1.1 | 66.4 | 4.81 | 93.95 | bef/dam |
| 7.3.1.2 | 80.1 | 1.23 | 48.4 | bef/djo |
| 7.3.1.3 | 84.4 | 2.43 | 60.2 | bef/squ |
| 7.3.1.4 | 58.6 | 3.4 | 38.6 | dam/djo |
| 7.3.1.5 | 73.5 | 2.05 | 67.9 | dam/squ |
| 7.3.1.6 | 85.6 | 2.83 | 99.7 | djo/squ |
| 7.3.2 | 2.82, 1.21 | 96.9, 55.8 | overall | |
| 7.4.1.1 | 73 | 0.581 | 68.4 | dam/sir |
| 7.4.1.2 | 77.4 | 3.25 | 53.7 | sav/san |
| 7.4.1.3 | 67.4 | 3.01 | 99.9 | sav/sob |
| 7.4.1.4 | 74.6 | 3.1 | 70.5 | sav/soub |
| 7.4.1.5 | 85.8 | 2.46 | 88.1 | sav/squ |
| 7.4.1.6 | 83.1 | 3.44 | 52.3 | sav/yah |
| 7.4.1.7 | 66.2 | 1.15 | 35.2 | san/sob |
| 7.4.1.8 | 68.8 | 1.15 | 93.2 | san/soub |
| 7.4.1.9 | 88.9 | 6.94 | 80.51 | san/squ |
| 7.4.1.10 | 77.3 | 1.74 | 98.1 | san/yah |
| 7.4.1.11 | 82.2 | 0.43 | 81.0 | sob/soub |
| 7.4.1.12 | 86.2 | 3.71 | 95.4 | sob/squ |
| 7.4.1.13 | 87.4 | 0.76 | 52.4 | sob/yah |
| 7.4.1.14 | 78.6 | 2.64 | 96.0 | soub/squ |
| 7.4.1.15 | 86.3 | 0.807 | 81.0 | soub/yah |
| 7.4.1.16 | 79.6 | 3.02 | 74.8 | squ/yah |
| 7.4.2 | 4.04, 0.97 | 76.9, 98.8 | overall | |
| 8.3.1.1 | 64.5 | 4.36 | 61.8 | sav/san |
| 8.3.1.2 | 71 | 3.84 | 89.9 | sav/soub |
| 8.3.1.3 | 76 | 2.01 | 69.7 | sav.squ |
| 8.3.1.4 | 82 | 3.54 | 59.1 | sav/yah |
| 8.3.1.5 | 73 | 0.73 | 79.3 | san/soub |
| 8.3.1.6 | 79.7 | 4.82 | 98.4 | san/squ |
| 8.3.1.7 | 77.3 | 1.44 | 98.5 | san/yah |
| 8.3.1.8 | 80.2 | 2.62 | 91.7 | soub/squ |
| 8.3.1.9 | 83.7 | 0.62 | 75 | soub/yah |
| 8.3.1.10 | 78.1 | 3.64 | 94.7 | squ/yah |
| 8.3.2 | 77.4 | 4.05, 0.87 | 87.4, 82.0 | overall |

¹Numbers refer to chapter headings

²Percentage of variation along the first principal component of the pooled within-groups correlation matrix

³Canonical root resulting from multivariate analysis of covariance with size as the covariable

⁴Ratio of the size-free canonical root to the usual canonical root expressed as a percentage.

⁵dam=*S. damnosum s.s.*, sir=*S. sirbanum*, san=*S. sanctipauli*, soub=*S. soubrense*, sob=*S. soubrense* 'B', bef=*S. soubrense* 'Beffa', djo=*S. sanctipauli* 'Djodji', squ=*S. squamosum*, yah=*S. yahense*, sav='savanna' (pooled *S. sirbanum* and *S. damnosum s.s.*)

9.3.3 COLOUR CHARACTERS

Two colour characters were included in this project (Chapter four) because both were considered to be of taxonomic importance in the *S. damnosum* complex, wing tuft colour (e.g. Garms 1978) and abdominal setal colour (Garms and Zillman 1984). The special nature of these characters was mentioned in Chapter seven, where a method was described which exploited the taxonomic potential of these characters without risking certain statistical assumptions being invalidated.

9.3.3.1 Wing Tuft Colour

Figure 9.1 gives the frequency histograms for each species combined as in Chapter eight. The taxonomic importance of this character is obvious from this figure, but it is also clear that there is considerable overlap between species for the five categories of the character.

The original five character state system for scoring this character of Kurtak *et al.* was used. This system probably obscures more subtle expression of this character, so it is recommended that in future studies, either more categories be defined (thus splitting the heterogeneous middle category), or the character be expressed differently (such as a proportion).

The overall 'global' discriminant analysis (Chapter eight) showed that, in practice, the influence of this character on helping allocation of flies to each of the five species was negligible. Whilst it is extremely useful for distinguishing between, for example, 'savanna' and *S. yahense*, these species are already well separated morphometrically, whereas those species which need extra information to aid discrimination, e.g. *S. soubrense* and *S. yahense*, overlap considerably for this character.

To conclude, the taxonomic importance of wing tuft colour has been confirmed in this project, although the method of scoring it has been criticised. In practise, the character helps to identify species which are already well distinguished, so its use is likely to be restricted to allocating flies of doubtful affinity after typicality probability.

9.3.3.2 Abdominal Setal Colour

The colour of the setae on the ninth abdominal tergite was recorded as a two-state character (Garms and Zillman 1984). In this project, it was found to be 95.8% diagnostic for *S. yahense*, meaning that the effect of this character on adjusted prior probabilities was diagnostic. The main problem with such extreme influence of a single character is that those *S. yahense* with character state one (white setae) are automatically incorrectly allocated, as are flies of other species with character state two, should they exist.

This extreme influence may be justified, as the effect of adjusting prior probabilities of species membership using this character is beneficial in discriminating between *S. yahense* and the species to which it is morphologically closest, *S. sanctipauli* and *S. soubrense*. However, the objectivity of scoring this character is not certain, and it is recommended that a detailed, double-blind study be undertaken to establish the true taxonomic status of this character.

9.3.4 COMMENTS ON DEPARTURES FROM EQUAL DISPERSION

The null hypothesis of equal dispersion was tested on each of the discriminant analyses because genuine rejection of this assumption is potentially serious, more so than departure from normality on the allocation rate of a given character subset (Campbell 1978). The

pooled dispersion matrix was used throughout the analyses because to calculate separate dispersion matrices for each group in an analysis involves calculating more parameters with less accuracy. The effect, generally would be to introduce optimistic bias into the estimate of error rate for a particular analysis.

In the between-species analyses given in Chapters seven and eight, the null hypothesis was tested 35 times and rejected 12 times at a significance level of $p < 0.0001$. Clearly, even though the testing of each null hypothesis is not independent, its rejection is too frequent to be by chance alone. A larger proportion of the rejections occurred in the 'global' analyses than in the regional analyses, which may be due to pooling resulting in departures from normality, to which the likelihood ratio test is particularly sensitive.

To conclude, given the poor performance of the statistical test for equal dispersion (Seber 1984), it is not possible to state with certainty whether the larger than expected rejections of the null hypotheses is due to genuine difference in covariance structure between species, or due to other causes. Determining the true nature of between-species variation in covariance structure is very important, and will need to be investigated in greater detail once more comprehensive sampling has been obtained. If differences in covariance structure exist, then allocation should be by the Quadratic discriminant function, or by typicality probability allocation to each species' mean vector weighted by the inverse of its own, rather than the pooled, dispersion matrix (Ambergen and Schaafsma 1984).

9.4 WORKED EXAMPLE OF ALLOCATION PROCEDURES

Two basic allocation procedures were used in these analyses, forced allocation and typicality allocation. Forced allocation assumes that the fly belongs to one of the reference species with probability 1, whereas typicality probability allocation calculates the probability that a fly is sampled from each species without reference to the other species. Forced allocation can be achieved in a number of ways, calculating the fly's score on each species' LDF, calculating each fly's posterior probability of species membership derived from its Mahalanobis' distance, and adjusting the prior probability of species membership in a way which reflects prior belief about the probability of the fly being one or other species. In this analysis, prior probabilities were calculated on the basis of colour, but other statements of prior belief could be used to adjust the probabilities.

As an example of the calculations of the statistics necessary for the different allocation procedures, three flies will be allocated using the overall 'global' statistics presented in section 8.3.1. Two (A and B) were collected from Bioko, West Africa by Dr. J. Mas of the Universidad de Barcelona. These flies belong to the *S. squamosum* subcomplex (i.e., *S. squamosum*, *S. yahense*, Post unpublished cytotaxonomic results). The other fly (C), was collected by Dr. P.J. McCall, Liverpool School of Tropical Medicine biting at cattle at Cynwyd, North Wales, and belonged to the species *S. variegatum*.

The vector of observations for the three flies, for the 13 character set given in Chapter 8 is shown below:

| Character | Fly | | |
|---------------------|---------|---------|---------|
| | A | B | C |
| Thorax Width | 968.95 | 993.792 | 1043.48 |
| Antennal Length 1 | 323.7 | 331.5 | 304.2 |
| Antennal Segment 4 | 47.12 | 42.16 | 44.64 |
| Antennal Segment 6 | 42.16 | 47.12 | 44.64 |
| Wing Length 1 | 777.36 | 787.2 | 1230.0 |
| Wing Length 2 | 492.0 | 492.0 | 747.84 |
| Wing Width 1 | 1029.35 | 1029.35 | 1499.91 |
| Wing Width 2 | 1529.32 | 1544.02 | 1970.47 |
| Wing Length 3 | 1646.96 | 1646.96 | 2088.11 |
| Femur Length 1 | 688.8 | 678.96 | 797.04 |
| Basitarsus Length 1 | 492.0 | 482.16 | 619.92 |
| Tibia Length 2 | 678.96 | 669.12 | 797.04 |
| Basitarsus Length 2 | 358.8 | 354.9 | 460.2 |

i). LDF Allocation

Computationally, this is the simplest method. The transposed vector of observations for each fly is postmultiplied by the vector part of each linear discriminant function given in Table 8.60, and the resultant score added to the constant. The three flies A, B and C score on each species' LDF in the following way,

| Species | Fly | | |
|-----------------------|--------|--------|--------|
| | A | B | C |
| savanna | 213.62 | 215.45 | 258.55 |
| <i>S. sanctipauli</i> | 217.16 | 220.88 | 241.29 |
| <i>S. soubrense</i> | 219.89 | 224.73 | 241.85 |
| <i>S. squamosum</i> | 218.07 | 219.1 | 272.56 |
| <i>S. yahense</i> | 223.13 | 225.96 | 255.38 |

Based on the highest score, both Bioko flies would be allocated into *S. yahense* and the *S. variegatum* would be allocated into *S. squamosum*.

Prior probabilities could be adjusted according to wing tuft and abdominal setal colour using the priors given in Tables 8.52 or 8.53 by adding $\ln \pi_i$ to each constant.

Although it is not necessary to calculate the discriminant functions each time they are to be used, this can be done if it is decided to write a general allocation program, rather than simply storing each species' LDF. The vector part of the LDF is calculated by $S^{-1}\bar{x}_i$ and the constant part by $\bar{x}_i'S^{-1}\bar{x}_i$.

ii). Posterior Probability Allocation

The first step in allocation is to calculate the Mahalanobis' squared distance to each of the five reference mean vectors shown in Table 8.58 using the inverse of the pooled within-species dispersion matrix shown in Table 8.59.

$$D_i^2 = (x-\bar{x})'S^{-1}(x-\bar{x})$$

These distances D_i^2 , are given below:

| Species | Fly | | |
|-----------------------|-------|-------|--------|
| | A | B | C |
| savanna | 37.34 | 37.4 | 342.32 |
| <i>S. sanctipauli</i> | 30.27 | 26.54 | 376.85 |
| <i>S. soubrense</i> | 24.81 | 18.85 | 375.72 |
| <i>S. squamosum</i> | 28.44 | 30.11 | 314.3 |
| <i>S. yahense</i> | 18.31 | 16.38 | 348.7 |

If forced allocation without adjusted prior probabilities is being used, then this is equivalent to assigning the fly to the species to which it is closest, thus the two Bioko flies would both be allocated into *S. yahense* and the *S. variegatum* would be allocated into *S. squamosum*.

To calculate the posterior probabilities of species membership, then the following calculation is performed:

$$R = \pi_i \exp(-0.5 D_i^2) / \text{SUM } \pi_i \exp(-0.5 D_i^2),$$

where summation is over $i=1\dots g$, the number of reference groups (5), and π_i is the prior probability of species membership taken either from Table 8.52 or Table 8.53.

The following table gives the posterior probabilities of species membership for each of the three flies, without adjusted prior probabilities:

| Species | Fly | | |
|-----------------------|--------|--------|---------|
| | A | B | C |
| savanna | <0.001 | <0.001 | <0.0001 |
| <i>S. sanctipauli</i> | 0.0024 | 0.0048 | <0.0001 |
| <i>S. soubrense</i> | 0.037 | 0.224 | <0.0001 |
| <i>S. squamosum</i> | 0.006 | 0.0008 | 0.999 |
| <i>S. yahense</i> | 0.954 | 0.7703 | <0.0001 |

Thus the posterior probabilities of species membership are highest for *S. yahense* for the two Bioko flies, and highest for *S. squamosum* for the *S. variegatum* fly.

If the prior probabilities are adjusted for wing tuft colour, then the probabilities in Table 8.52 are used. Fly A has wing tuft colour category 4 and therefore enters the table at the fourth row, fly B has colour category 5 and enters the table in the final row, while fly C has colour category 1 and enters the table in row 1.

The following table gives the posterior probability of species membership for each of the flies, with each fly's prior probability adjusted using the appropriate prior probabilities:

| Species | Fly | | |
|-----------------------|--------|---------|---------|
| | A | B | C |
| savanna | 0.0 | 0.0 | <0.0001 |
| <i>S. sanctipauli</i> | 0.0036 | <0.0001 | <0.0001 |
| <i>S. soubrense</i> | 0.028 | 0.018 | <0.0001 |
| <i>S. squamosum</i> | 0.0002 | <0.0001 | 0.999 |
| <i>S. yahense</i> | 0.9684 | 0.979 | <0.0001 |

Once again the three flies are allocated into *S. yahense*, *S. yahense* and *S. squamosum* respectively.

Adjusting the prior probabilities of species membership using abdominal setal colour and wing tuft colour involves using the prior probabilities shown in Table 8.53. The two Bioko flies had character state 2 for this character and so enter the table in the last row, and Fly C had character state one for both characters and so enters the table in the first row. The following table gives the posterior probabilities of species membership for each fly:

| Species | Fly | | |
|-----------------------|-----|-----|---------|
| | A | B | C |
| savanna | 0.0 | 0.0 | <0.0001 |
| <i>S. sanctipauli</i> | 0.0 | 0.0 | <0.0001 |
| <i>S. soubrense</i> | 0.0 | 0.0 | <0.0001 |
| <i>S. squamosum</i> | 0.0 | 0.0 | 0.999 |
| <i>S. yahense</i> | 1.0 | 1.0 | 0.0 |

Once again, this shows that the two Bioko flies show a strong affinity for *S. yahense* whilst, apparently the *S. variegatum* shows strong affinity for *S. squamosum*.

iii). Typicality Probability Allocation

However, to obtain greater detail of the affinities of the three flies for each of the reference species, it is necessary to calculate the typicality probabilities of species membership. Using the method

of Ambergen and Schaafsma (1984) first involves calculating an unbiased estimate of Mahalanobis' distance:

$$((n-g-p-1)/n)D_i^2 - p/n_i$$

where n = total sample size = 800,

g = number of reference groups = 5,

p = number of characters = 13,

n_i = sample size of i -th reference species ($i=1...5$).

The following table gives the unbiased distances of the three flies to each of the reference species:

| Species | Fly | | |
|-----------------------|-------|-------|--------|
| | A | B | C |
| savanna | 36.4 | 36.5 | 334.14 |
| <i>S. sanctipauli</i> | 29.33 | 25.69 | 367.69 |
| <i>S. soubrense</i> | 24.16 | 18.35 | 366.73 |
| <i>S. squamosum</i> | 27.69 | 29.32 | 306.78 |
| <i>S. yahense</i> | 17.74 | 15.85 | 340.23 |

Using these distances an estimate of its variance is then calculated:

$$(n-g-p-3)^{-1} \{2D^4 + 4(n-g-1)n_i^{-1} D^2 + 2p(n-g-1)n_i^{-2}\}$$

where n is the total sample size on which the dispersion matrix is based, g is the number of reference species, p is the number of characters and n_i is the sample size of the i -th species.

The unbiased distance plus and minus half the variance gives an approximate confidence interval for the distance of the fly to each of the reference species:

| Species | Fly | | |
|-----------------------|------------|------------|---------------|
| | A | B | C |
| savanna | 36.4±2 | 36.5±2 | 334.14±146.1 |
| <i>S. sanctipauli</i> | 29.33±2.1 | 25.69±1.7 | 367.69±185.8 |
| <i>S. soubrense</i> | 24.16±0.97 | 18.35±0.6 | 366.73±176.0 |
| <i>S. squamosum</i> | 27.69±1.3 | 29.32±1.45 | 306.78±124.32 |
| <i>S. yahense</i> | 17.74±0.78 | 15.85±0.66 | 340.23±155.8 |

The upper and lower distances thus derived can then be referred to the χ^2 distribution with p-degrees of freedom (13), to give the approximate confidence intervals:

| Species | Fly | | |
|-----------------------|---------------|---------------|---------|
| | A | B | C |
| savanna | 0.001,0.0002 | 0.001,0.0002 | 0.0,0.0 |
| <i>S. sanctipauli</i> | 0.0115,0.0029 | 0.0313,0.011 | 0.0,0.0 |
| <i>S. soubrense</i> | 0.0395,0.0222 | 0.1673,0.1247 | 0.0,0.0 |
| <i>S. squamosum</i> | 0.0151,0.0065 | 0.0095,0.0036 | 0.0,0.0 |
| <i>S. yahense</i> | 0.2011,0.1386 | 0.2955,0.2225 | 0.0,0.0 |

Clearly, therefore, the *S. variegatum* shows no affinity for any of the reference species. This example is extreme, in that the Mahalanobis' distances of this fly to the five species were all very large, but the principle is clearly demonstrated: that forced allocation can sometimes lead to an unrelated or atypical fly being forced into a species to which it does not belong.

Flies A and B both show greater affinity for *S. yahense* than for any other species, and their confidence intervals do not overlap with any other species. Therefore, especially when considering the forced allocation results, both flies should be allocated to *S. yahense*.

9.5 FIELD PROTOCOL

The exact statistics to be used in identifying adult females of the *S. damnosum* complex in the field clearly depends on the specific aims of the individual project. The purpose of this section is to give a key to the analyses presented in Chapters seven and eight, and hence to the characters to be measured. Some suggestions will also be made for adjusting the methods to suit a specific situation which may arise in the field.

Key to Discriminant Analyses, West Africa.

- 1). Geographical region taken into account.....2
 No account taken of geography.....5
- 2). Togo and Benin.....3
 Western Area.....4
- 3). Overall discrimination only.....7.3.2
 Overall and species-pair discrimination.....7.3.1,7.3.2
 Species-pair discrimination only.....7.3.1.1-7.3.1.6
- 4). Overall discrimination only.....7.4.2
 Overall and species-pair discrimination.....7.4.1,7.4.2
 Species-pair discrimination only.....7.4.1.1-7.4.1.16
- 5). Overall discrimination only.....8.3.2
 Overall and species-pair discrimination.....8.3.1,8.3.2
 Species-pair discrimination only.....8.3.1.1-8.3.1.10

Examining the statistics in each section and referring to Table 9.1 will give the characters to be measured for each analysis. In addition, it is recommended that wing tuft colour and abdominal setal colour also be recorded.

There are advantages and disadvantages to each of the approaches which could be taken to identifying adult females using these statistics. The principal advantage of using the overall 'global' statistics to the species-pair 'global' statistics in Chapter eight is that sample sizes are larger, and so estimates of the parameters more accurate than the regional estimates. However, unlike samples are pooled, which may give unpredictable results in allocation. The advantage of the regional statistics is that only flies native to a region were used to derive the statistics for allocation in that region, but the disadvantage is that smaller sample sizes were used with consequent loss of resolving power.

The species-pair statistics have the advantage that the statistics were developed for just that species pair, but therefore the estimate of the pooled dispersion matrix is based on a smaller sample size than for the overall statistics. Also, if the assumption implicit in using only species-pair statistics is violated, that flies in an area only belong to those species, then results might be unpredictable. This does not occur if the species-pair statistics are used subsequent to overall allocation, but then there is the disadvantage that more characters need to be measured if all species-pair combinations are accounted for.

This last problem can be alleviated if certain of the species-pair statistics are discounted prior to the analysis, for example, if all combinations in section 7.4 were covered, then 22 characters would need to be measured on each fly as well as the two colour characters. However, if those analyses with size-free canonical roots less than 1.0 are discounted, then the number of characters falls to 16, plus the two colour characters. Likewise, the 'global' statistics require

23 characters, plus the two colour characters, but if species-pair analyses with size-free canonical roots less than 1.0 are discounted then this number falls to 18, plus the two colour characters.

In practise, in the field, it is unlikely that all of the species-pair statistics will be required, and these are more likely to be used in a laboratory where more sophisticated computers will be able to access the relevant statistics in a straightforward manner not likely to be practicable with the programmable calculator or portable computer envisaged as the field tool for this method. It is more likely that the overall statistics, whether 'global' or regional, together with certain species-pair statistics known to be of use in that area will be the only statistics used in the field method, therefore reducing the number of characters which need to be measured.

The statistics developed in Chapters seven and eight can be modified to meet the particular requirements of an area. For example, if it is known with some certainty from other evidence that it is extremely unlikely that 'savanna' flies will be found in an area, then the prior probabilities of species membership can be adjusted accordingly. If the 'global' overall statistics of Chapter eight are used, then the priors might be set at

0.0, 0.25, 0.25, 0.25, 0.25

for the five species 'savanna', *S. sanctipauli*, *S. soubrense*, *S. squamosum*, *S. yahense*. If the information in the wing tuft colour is also required, then this can simply be achieved by first multiplying the new priors and the appropriate colour character priors, summing the intermediate result, and dividing each intermediate result by this sum. For example, the first row of Table 8.52 is,

0.5644, 0.0009, 0.032, 0.4026, 0.0

multiplying this row by the new set of priors gives the intermediate result,

0.0, 0.000225, 0.008, 0.10065, 0.0

which sums to 0.1089. Dividing each of these by the sum gives,

0.0, 0.002, 0.0735, 0.9245, 0.0

the new set of prior probabilities. This would be done for each row of this table (except for the final one where the result is identical).

It must be emphasised, however, that objective justifications for altering the prior probability of a species must be established before this step can be taken.

A user of the statistics developed in this project may decide, on objective grounds that certain of the characters derived in any one of the discriminant analyses are unnecessary. If this is so, then the corresponding row and column of the pooled dispersion matrix and the relevant row of the mean vectors can simply be deleted, and the LDF also recalculated.

9.6 GENERAL CONCLUSIONS

The methods developed in this project for the identification of adult females of the *S. damnosum* complex in West Africa will clearly be of benefit to studies aimed at determining the relative vectorial importance of the different species in transmitting *O. volvulus*. The very successful identification of the 'savanna' species *S. damnosum* s.s. and *S. sirbanum* from all other species is of most benefit, as these species transmit the more debilitating strain of onchocerciasis. The ability to identify *S. squamosum* and *S. yahense* individually provides a more practical field identification method than enzyme electrophoresis. But the proximity of *S. sanctipauli*, *S. soubrense* and *S. yahense* means that unequivocal identification of these species cannot be established in the absence of abdominal setal colour, but the statistics presented in this project represent the best available method for distinguishing between them.

Of the two methods of allocation, forced allocation and typicality probability allocation, it is recommended that both be used simultaneously, as both can easily be written into the same computer program. Typicality probability allocation allows a more realistic assessment of a fly's affinities to be made, while forced allocation allows the inclusion of additional prior knowledge to influence the specific identity of a fly.

The multivariate statistical method for identification of adult females as it is presented in this project could very easily be applied as a field method, without requiring any laboratory facilities, and after very short training of field workers. This facility is a great improvement over other methods of adult identification, with the exception of current morphological methods.

There are still outstanding problems within the morphology and morphometrics of the *S. damnosum* complex which will need to be investigated in further detail, although the method presented in this project can still be used in their absence.

1). Wider geographic and seasonal variation needs to be sampled for each species in the *S. damnosum* complex, with the ultimate aim of providing discriminant statistics for specific geographic regions and climatic seasons.

2). More intensive sampling of chromosomally known samples of new cytoforms needs to occur, such as *S. sanctipauli* 'Djodji', the OP-insecticide resistant *S. sanctipauli*, and the various forms of *S. soubrense*. The morphological status of *S. dieguerense* needs to be established in light of recent data on more extensive geographic range of this species (Boakye and Mosha 1988).

A more sensitive system for scoring wing tuft colour should be adopted which more accurately represents the variation of this character.

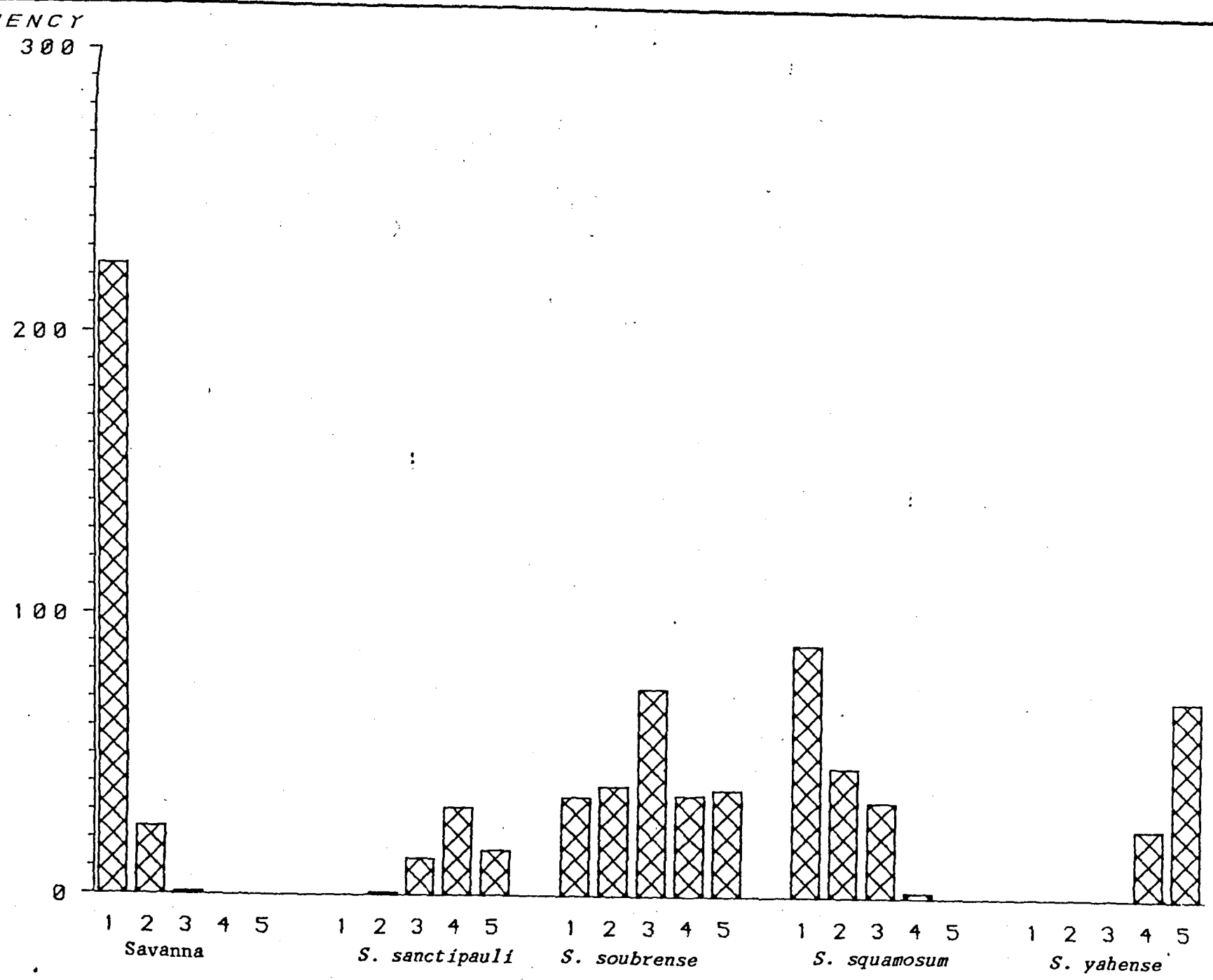
4). The reliability of abdominal setal colour as a taxonomic character for *S. yahense* (Garms and Zillman 1984) needs to be established using a double-blind trial.

5). A continuously updated data base should be established on computer of measurements on flies of known chromosomal identity containing the characters used in this project, those previously used in morphological studies of the *S. damnosum* complex, and any new characters as they are discovered. This will provide more reliable estimates of parameters such as mean vectors and dispersion matrix, and also allow an objective assessment of the error rates of the different morphological methods. The question of equal dispersion

will also be more easily established once larger sample sizes of each species are available.

Figure 9.1

Histograms of Wing Tuft Colour for *S. damnosum* s.l.



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APPENDIX ONE: DETAILS OF ADULT FEMALE SAMPLES USED FOR MULTIVARIATE
MORPHOMETRICS

Sample V1=1.

River Amou at Amou Oblo, Togo, 7° 24'N 0° 53'E, collected by Dr. R.J.

Post, 19/10/1984, reared from pupae, preserved in 95% propanol.

Species identity, *S. squamosum*, from correlated larval cytotaxonomy,

36 *S. squamosum*, determined by D.P. Surtees.

Initial sample size = 40

Number rejected because of missing values = 2

Number rejected as outliers = 4

Final sample size = 34

Sample V1=2.

River Nanie at Nigbi, Côte d'Ivoire, 5° 38'N 6° 38'W, collected by

Dr. R.J. Post, 10/07/1985, reared from pupae, preserved in 95%
propanol.

Species identity, *S. yahense*, from correlated larval cytotaxonomy,

109 *S. yahense*, (determined by D.P. Surtees).

Initial sample size = 41

Number rejected because of missing values = 2

Number rejected as outliers = 2

Final sample size = 37

Sample V1=3.

River Niger at Tienfala, Mali, 12° 43'N 7° 44'W, collected by Dr. R.J.

Post, 02/11/1984, reared from pupae, preserved in 95% propanol.

Species identity, *S. sirbanum*, from correlated larval cytotaxonomy,

7 *S. sirbanum*, determined by D.P. Surtees.

Initial sample size = 35

Number rejected because of missing values = 2

Number rejected as outliers = 4

Final sample size = 29

Sample V1=4.

River Sassandra at Soubre, Côte d'Ivoire, 5° 47'N 6° 37'W, collected

by Dr. R.J. Post, 26/10/1984, reared from pupae, preserved in
95% propanol.

Species identity, *S. sanctipauli*, from correlated larval
cytotaxonomy, 29 *S. sanctipauli*, determined by D.P. Surtees.

Initial sample size = 40

Number rejected because of missing values = 4

Number rejected as outliers = 1

Final sample size = 35

Sample V1=5.

River Gban-Houa at Djodji, Togo, 7° 42'N 0° 35'E, collected by Dr. R.J. Post, 15/10/1984, reared from pupae, preserved in 95% propanol.

Species identity, *S. sanctipauli* 'Djodji'/*S. squamosum*, from correlated larval cytotaxonomy, 23 *S. sanctipauli* 'Djodji': 11 *S. squamosum*, determined by D.P. Surtees. Mixture separated by external analysis, using LDFs derived from samples V1=1 and V1=4, and internal analysis using PCA and cluster analysis.

Initial sample size = 50

Number rejected because of missing values = 4

Number rejected as outliers (following separation of mixture) = 1

Final sample size, V1=1 (*S. sanctipauli* 'Djodji') = 26 Final sample size, V1=30 (*S. squamosum*) = 17

Sample V1=6.

River Bankasoka at Port Loko, Sierra Leone, 8° 45'N 12° 47'W, collected by MRC laboratory staff, Bo, 25/01/1986, reared from pupae, preserved in 95% propanol.

Species identity, *S. soubrense* 'B', from previous chromosomal identification (Post 1986).

Initial sample size = 37

Number rejected because of missing values = 7

Number rejected as outliers = 2

Final sample size = 28

Sample V1=7.

River Waanje at Kenema Waterfall, Sierra Leone, 7° 54'N 11° 14'W,
collected by Dr. R.J. Post, 21/07/1985, reared from pupae, pre-
served in 95% propanol.

Species identity, *S. yahense*, from correlated larval cytotaxonomy,
38 *S. yahense*, determined by D.P. Surtees.

Initial sample size = 20

Number rejected because of missing values = 4

Number rejected as outliers = 2

Final sample size = 14

Sample V1=8.

River Bebeye at Gerihun, Sierra Leone, 7° 57'N 11° 35'W, collected
by D.P. Surtees, Dr. J.B. Davies, M.C. Thomson, 20/01/1986,
reared from pupae, preserved in 95% propanol.

Species identity, *S. yahense*, from previous chromosomal identifica-
tions (Post 1986).

Initial sample size = 40

Number rejected because of missing values = 12

Number rejected as outliers = 3

Final sample size = 25

Sample VI=9.

River Taia at Mongeri, Sierra Leone, 8° 19'N 11° 44'W, collected by D.P. Surtees, Dr. J.B. Davies, M.C. Thomson, 23/01/1986, reared from pupae, preserved in 95% propanol.

Species identity unknown.

Initial sample size = 7

Sample not used in subsequent analyses.

Sample VI=10.

River Seli at Yirafilaia, Sierra Leone, 9° 28'N 11° 20'W, collected by D.P. Surtees, Dr. J.B. Davies, M.C. Thomson, 09/02/1986, reared from pupae, preserved in 95% propanol.

Species identity, *S. damnosum* s.s.: (*S. squamosum*/*S. yahense*), from previous larval cytotoxic identification (4 *S. damnosum* s.s., 7 *S. squamosum*) determined by D.P. Surtees. Also, correlated DNA probes identifications of same adults (Post pers. comm.), 27 *S. damnosum* s.s., 6 *S. squamosum*/*S. yahense*, 3 unknown.

Initial sample size = 40

Number rejected because of missing values = 4

Number rejected as outliers/ contaminants = 8

Final sample size = 28

Sample V1=11.

River Kakatemadaru, Benin, 10° 07'N 03° 20'E, collected by Dr. R.A. Cheke, 10/09/1984, caught at human bait, preserved in 70% ethanol.

Species identity, *S. damnosum* s.s., Dr. R.A. Cheke personal communication.

Initial sample size = 20

Number rejected because of missing values = 4

Number rejected as outliers = 0

Final sample size = 16

Sample V1=12.

River Amou at Amou Oblo, Togo, 7° 24'N 0° 53'E, collected by Dr. R.A. Cheke, 14/03/1985, reared from pupae, preserved in 95% propanol.

Species identity, *S. squamosum*, from correlated larval cytotaxonomy, 40 *S. squamosum* determined by D.P. Surtees.

Initial sample size = 18

Number rejected because of missing values = 0

Number rejected as outliers = 4

Final sample size = 14

Sample VI=13.

River Amoutchou at Idifiou, Togo, 7° 38'N 0° 58'E, collected by Dr. R.A. Cheke, A.M. Denke, 09/10/1985, reared from pupae, preserved in 95% propanol.

Species identity, *S. squamosum*, from correlated larval cytotaxonomy, 30 *S. squamosum* determined by D.P. Surtees.

Initial sample size = 29

Number rejected because of missing values = 6

Number rejected as outliers = 3

Final sample size = 20

Sample VI=14.

River Mono at T52, Togo, 6° 54'N 01° 36'E, collected by Dr. R.A. Cheke, 02/11/1981, caught bitong on man, preserved in 70% ethanol.

Species identity, *S. soubrense* 'Beffa', Dr. R.A. Cheke personal communication.

Initial sample size = 32

Number rejected because of missing values = 5

Number rejected as outliers = 2

Final sample size = 25

Sample V1=15.

River Baoule at Wandadou, Guinea, 9° 04'N 09° 20'W, collected by Dr.

R. Baker, Sept. 1986, caught at human bait, preserved in 95%
propanol.

Species identity, *S. squamosum*, Dr. R. Baker personal communication.

Also, correlated DNA probes identifications, 43 *S. squamosum*,
Dr. R.J. Post personal communication.

Initial sample size = 45

Number rejected because of missing values = 1

Number rejected as outliers = 4

Final sample size = 40

Sample V1=16.

River Milo at Balan, Guinea, 9° 46'N 09° 10'W, collected by Dr. R.

Baker, Sept. 1986, caught at human bait, preserved in 95%
propanol.

Species identity, *S. soubrense*, Dr. R. Baker personal communication.

Initial sample size = 39

Number rejected because of missing values = 1

Number rejected as outliers = 8

Final sample size = 30

Sample V1=17.

River Milo at Konsankoro, Guinea, $9^{\circ} 02'N 09^{\circ} 00'W$, collected by Dr.

R. Baker, Sept. 1986, caught at human bait, preserved in 95%
propanol.

Species identity, *S. soubrense*, Dr. R. Baker personal communication.

Initial sample size = 38

Number rejected because of missing values = 2

Number rejected as outliers = 4

Final sample size = 32

Sample V1=18.

River Makona at Yalamba, Guinea, $8^{\circ} 31'N 10^{\circ} 11'W$, collected by Dr.

R. Baker, Sept. 1986, caught at human bait, preserved in 95%
propanol.

Species identity, *S. soubrense*, Dr. R. Baker personal communication.

Initial sample size = 40

Number rejected because of missing values = 1

Number rejected as outliers = 1

Final sample size = 38

Sample V1=19.

River Bafing at Koukoutamba, Guinea, 11° 17'N 11° 20'W, collected by
Dr. R. Baker, Sept. 1986, caught at human bait, preserved in 95%
propanol.

Species identity, *S. soubrense*, Dr. R. Baker personal communication.

Initial sample size = 38

Number rejected because of missing values = 1

Number rejected as outliers = 6

Final sample size = 31

Sample V1=20.

River Koudeta at Bassi, Guinea, 10° 51'N 11° 14'W, collected by Dr.
R. Baker, Sept. 1986, caught at human bait, preserved in 95%
propanol.

Species identity, *S. squamosum*, Dr. R. Baker personal communication.

Initial sample size = 35

Number rejected because of missing values = 0

Number rejected as outliers = 1

Final sample size = 34

Sample V1=21.

River Niger at Laya Doula, Guinea, $09^{\circ} 51'N 10^{\circ} 39'W$, collected by
Dr. R. Baker, Sept. 1986, caught at human bait, preserved in 95%
propanol.

Species identity, *S. sirbanum*, Dr. R. Baker personal communication.

Initial sample size = 45

Number rejected because of missing values = 1

Number rejected as outliers = 0

Final sample size = 44

Sample V1=22.

River Bouka 2 at Sidakele, Guinea, $11^{\circ} 30'N 10^{\circ} 10'W$, collected by
Dr. R. Baker, Sept. 1986, caught at human bait, preserved in 95%
propanol.

Species identity, *S. sirbanum*, Dr. R. Baker personal communication.

Initial sample size = 33

Number rejected because of missing values = 1

Number rejected as outliers = 3

Final sample size = 29

Sample V1=23.

River Mafou at Serekoroba, Guinea, 10° 24'N 10° 09'W, collected by
Dr. R. Baker, Aug. 1985, caught at human bait, preserved in 70%
ethanol.

Species identity, *S. sirbanum*, Dr. R. Baker personal communication.

Initial sample size = 50

Number rejected because of missing values = 1

Number rejected as outliers = 6

Final sample size = 43

Sample V1=24.

River Niger at Diaragbela, Guinea, 10° 36'N 09° 59'W, collected by
Dr. R. Baker, Dec. 1985, caught at human bait, preserved in 70%
ethanol.

Species identity, *S. sirbanum*, Dr. R. Baker personal communication.

Initial sample size = 35

Number rejected because of missing values = 0

Number rejected as outliers = 2

Final sample size = 33

Sample V1=25.

River Bale at Menankaya, Guinea, 09° 33'N 09° 35'W, collected by Dr. R. Baker, Aug. 1985, caught at human bait, preserved in 70% ethanol.

Species identity, *S. soubrense*, Dr. R. Baker personal communication.

Initial sample size = 27

Number rejected because of missing values = 0

Number rejected as outliers = 0

Final sample size = 27

Sample V1=26.

River Niger at Mamouria, Guinea, 09° 23'N 10° 34'W, collected by Dr. R. Baker, Sept. 1986, caught at human bait, preserved in 70% ethanol.

Species identity, *S. squamosum/S. yahense*, Dr. R. Baker personal communication.

Initial sample size = 20

Number rejected because of missing values = 0

Number rejected as outliers = 7

Final sample size = 13

Sample V1=27.

River Makona at Bofossou, Guinea, 08° 39'N 09° 41'W, collected by Dr. R. Baker, Sept. 1986, caught at human bait, preserved in 70% ethanol.

Species identity, *S. yahense*/*S. squamosum*, Dr. R. Baker personal communication.

Initial sample size = 25

Number rejected because of missing values = 1

Number rejected as outliers = 4

Final sample size = 20

Sample V1=28.

River Anie at Konogbe, Togo, 7° 48'N 1° 5'E, collected by OCP insecticide team, 12/10/84, reared from pupae, preserved in 95% propanol. Species identity, *S. damnosum s.s.* and *S. soubrense* 'Beffa' mix, 14 *S. damnosum s.s.*, 4 *S. soubrense* 'Beffa' determined by D.P. Surtees. Mixture separated by external analysis, using LDFs, and internal analysis using PCA and cluster analysis.

Initial sample size = 43

Number rejected because of missing values = 2

Number rejected as outliers (following separation of mixture) = 4

Final sample size, V1=28 (*S. damnosum s.s.*) = 27 Final sample size, V1=29 (*S. soubrense* 'Beffa') = 11

Sample V1=29.

See V1=28.

Sample V1=30.

See V1=5.

The cytotaxonomy of the Djodji form of *Simulium sanctipauli* (Diptera: Simuliidae)

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Summary

The Djodji form is described as a new cytotype of *Simulium sanctipauli* within the *S. damnosum* complex, from the Ghana/Togo border area on the basis of sex chromosome differentiation.

Introduction

The potential importance of describing genetically distinct forms or geographic races within previously recognised cyto-species of vector complexes such as the *Simulium damnosum* complex comes from the possible correlation of the different forms with factors of epidemiological significance, such as anthropophily, together with their usefulness in tracing migration patterns or insecticide resistance distribution.

The purpose of this paper is to describe a new cytotoxic form within *S. sanctipauli* from Ghana and Togo, and to give criteria for its routine identification in any future studies of onchocerciasis transmission in that area.

Materials and methods

Breeding sites where the Djodji form of *S. sanctipauli* was collected are listed in Table 1. Larvae were fixed in 3:1 ethanol:acetic acid and stored in a refrigerator. For preparation of polytene chromosomes the larvae were split open ventrally and hydrolysed in hydrochloric acid. The silk glands were stained in feulgen and/or orcein following standard methods (Vajime and Dunbar 1975, Quillévert 1975, Post 1986). Larval sex was determined according to the shape of the developing gonads (Puri 1925) after staining with feulgen. Inversions were scored from polytene chromosome preparations by comparison with the standard maps of Post (1986).

Chromosomal characteristics and cytotoxic key

All fixed and polymorphic inversions within Djodji form are indicated on the idiogram (Fig. 1), and frequencies of polymorphic inversions are listed in Table 2. The new form is homozygous for the fixed inversions 1L-P&Q, 2L-4&6&A and 3L-2, but there are no fixed inversions unique to Djodji form, and only one new rare polymorphic inversion (1S-P, see Figure 2). The presence of inversion 2L-A places Djodji form within *S.*

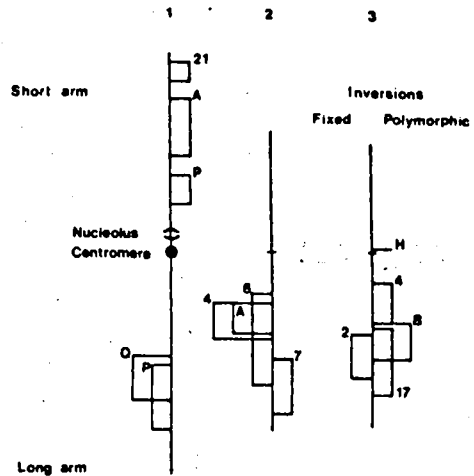


Fig. 1 Idiogram showing the relative positions of the breakpoints of all the inversions currently known from the Djodji form of *S. sanctipauli*. Those inversions plotted to the right are intraspecific polymorphisms, whilst those to the left are fixed inversions plotted on the idiogram to indicate the derivation of the Standard sequence (as seen in *S. squamosum*) from the basic Djodji sequence. The polymorphic inversion 3L-B is based on the 3L-4.17.2 sequence, and not the 3L-2 sequence. Most of these inversions are illustrated by Post (1986), although 1S-21, 1S-A and the new inversion 1S-P are also shown in Figures 2 and 3

sanctipauli (Post 1986). However, 1S-21 (Figure 3) is strongly Y-linked in Djodji form (Table 2), and this unique feature is the most important cytotoxic criterion for both description and routine identification.

Since 1S-21 is Y-linked in Djodji form there is no single inversion which is diagnostic of all individuals. However, samples in which there is strong Y-linkage of the inversion can be unequivocally identified as *S. sanctipauli* Djodji form, and mixed samples (should they exist) of the form with typical *S. sanctipauli*, will be recognised as such using standard population genetic analysis.

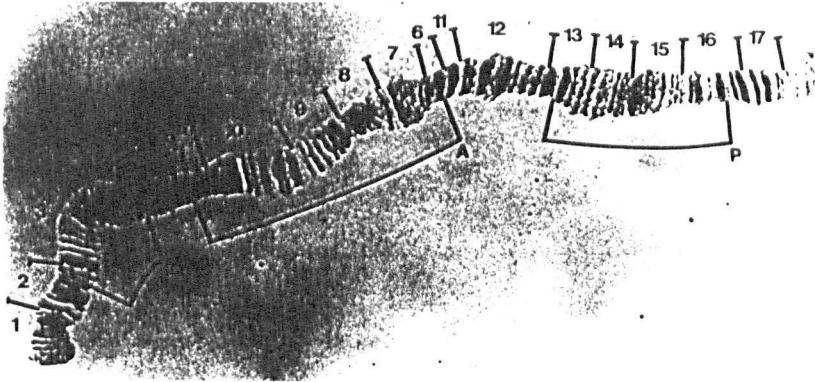


Fig. 2 *S. sanctipauli* (collected by R. J. Post from the river Sassandra at Soubré 26.10.84) showing the karyotype 1S-A/A with the breakpoints of inversions 1S-21 and 1S-P indicated

Table 1 List of larval samples from which the Djodji form of *S. sanctipauli* has been identified

Code: River; collection number, co-ordinates (N/E), date, cytospecies composition¹, collectors².

¹sq = *S. squamosum*, ya = *S. yahense*, sa = Djodji form of *S. sanctipauli*, da = *S. damnosum* s.s. and si = *S. sirbanum*. Other cytospecies were not found in these samples.

²AKA = A. K. Adzah, AKO = A. K. Opoku, AMD = A. M. Denke, CKP = C. P. Kowal Post, EAW = E. A. Weber, HSA = H.S.K. Avissey, JEH = J. E. E. Henderickx, JFW = J. F. Walsh, MA = M. Ampah, MD = M. David, RAC = R. A. Cheke, RJP = R. J. Post, SS = S. Sowah, SSH = OCP subsector Hohoe, YY = Y. Yamagata.

Dayi: 1, 7°09'0"29', 13.01.87, 22sq 5ya 3sa 1da, AKA. 2, 7°07'0"27', 22.01.87, 20sq 74ya 8sa 1da, RAC EAW YY. 3, 7°06'0"26', 17.03.86, 31sq 12ya 4sa AKA. 4, 7°06'0"26', 29.04.86, 20sq 5ya 4sa, SSH. 5, 6°57'0"21', 12.02.86, 17sq 15ya 21sa 55da 2si, AKA. 6, 6°57'0"21', 21.03.86, 9sq 8sa 47da AKA. 7, 6°53'0"21', 06.05.86, 2sq 4sa 8da 3si, SSH. 8, 6°53'0"21', 16.05.86, 5sq 2ya 17sa 16 da 8si, SSH. 9, 6°53'0"21', 13.01.87, 2sq 4sa 8da 3si, RAC EAW YY. 10, 6°52'0"19', 29.05.86, 2sq 8sa 19da 2si, AKA. 11, 6°52'0"19', 22.01.87, 2sq 1sa 7da, RAC EAW YY.

Asukawkaw: 12, 7°54'0"37', 05.02.87, 54sq 7sa JFW JEEH. 13, 7°54'0"37', 27.05.86, 41sq 13sa, YY. 14, 7°54'0"37', 23.01.87, 91sq 27sa, RAC EAW YY. 15, 7°52'0"36', 18.03.86, 26sq 4sa, MD. 16, 7°52'0"29', 23.01.87, 33sq 28sa, RAC EAW YY. 17, 7°41'0"26', 28.05.86, 29sq 52sa, YY. 18, 7°41'0"25', 06.02.86, 5sq 90sa. 19, 7°41'0"25', 23.01.87, 2sq 74sa, RAC EAW YY.

Menou: 20, 7°37'0"39', 28.05.86, 28sq 14sa, YY.

Gban-Houa: 21, 7°42'0"38', 28.05.86, 3sq 25sa, YY. 22, 7°42'0"38', 23.01.87, 12sq 29sa, RAC EAW YY. 23, 7°41'0"37', 06.02.86, 12sq 18sa, YY. 24, 7°42'0"36', 23.01.87, 11sq 22sa 1da, RAC EAW YY. 25, 7°42'0"35', 15.10.84, 4sq 29sa, RJP CKP. 26, 7°42'0"35', 15.03.85, 17sq 23sa, RAC AMD. 27, 7°42'0"35', 21.03.85, 11sq 23sa, RAC AMD. 28, 7°42'0"35', 26.03.85, 8sq 32sa, RAC AMD. 29, 7°42'0"35', 29.03.85, 10sq 29 sa, RAC AMD. 30, 7°42'0"35', 15.10.85, 13sq 21sa, RAC AMD. 31, 7°42'0"35', 15.03.86, 45sq 22sa, YY. 32, 7°42'0"35', 27.01.87, 28sq 54sa 1da, RAC HSA.

Wawa: 33, 7°43'0"33', 20.03.86, 13sq 12sa, YY. 34, 7°43'0"33', 23.01.87, 13sq 32sa, RAC EAW YY. 35, 7°41'0"30', 27.03.86, 17sq 6sa, AKA.

Kpaza: 36, 8°33'0"41', 15.10.87, JFW SS. 37, 8°33'0"41', 22.10.87, JFW YY AKO. 38, 8°33'0"37', 22.10.87, JFW YY AKO.

Niankpe: 39, 9°05'0"42', 21.10.87, JFW YY AKO.

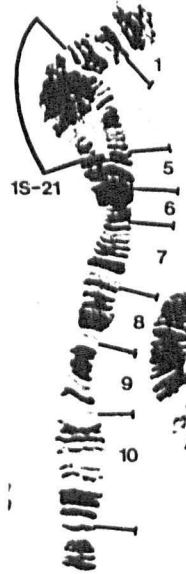


Fig. 3 The Djodji form of *S. sanctipauli* from the river Gban-Houa at Djodji showing the karyotype 1S-St/21

The following cytotoxic key can be used for the identification of *S. soubrense*, *S. soubrense* Beffa form, *S. sanctipauli* and the Djodji form of *S. sanctipauli*. The key should not be used west of Côte d'Ivoire, where *S. soubrense* form Konkouré and *S. soubrense* 'B' might also be encountered.

- 1) Larva homozygous for inversions 1L-P&Q, 2L-4&6 and 3L-2
 - *S. sanctipauli* subcomplex 2)
 - These inversions absent from larva
 - Other species of *S. damnosum* complex

Table 2 Inversion frequencies in the Djodji form of *S. sanctipauli*

| River | Samples ¹ | IS-21 karyotype frequency | | | | | | Autosomal polymorphic inversion frequencies ² | | | | | |
|----------------|---|---------------------------|-------|-------|--------------------|-------|-------|--|------|------|---------|-------|-------|
| | | Numbers of males | | | Numbers of females | | | IS-A | 2L-7 | 3L-B | 3L-4.17 | 3L-24 | IS-P |
| | | st/st | st/21 | 21/21 | st/st | st/21 | 21/21 | | | | | | |
| Dayi | 1, 2, 4, 7 8, 9, 11 | 2 | 14 | 0 | 22 | 1 | 0 | 0.56 | 1.00 | 0 | 1.00 | 0 | 0 |
| Asukaw- kaw | 14, 16, 19 | 0 | 58 | 0 | 31 | 2 | 0 | 0.42 | 1.00 | 0 | 0.97 | 0 | 0 |
| Gban- Houa | 22, 24, 25 26, 27, 28, 29, 30, 32 | .3 | 108 | 0 | 113 | 1 | 0 | 0.45 | 1.00 | 0 | 0.996 | 0.002 | 0.002 |
| Wewa | 34 | 1 | 12 | 0 | 8 | 0 | 0 | 0.50 | 1.00 | 0 | 1.00 | 0 | 0 |
| Kpaza | 36, 37 | 0 | 10 | 0 | 11 | 0 | 0 | 0.31 | 1.00 | 0 | 1.00 | 0 | 0 |

¹Samples are as listed in Table 1. However, it was not always possible to score all inversions in every specimen, and hence sample sizes may be slightly smaller than those listed in Table 1.

²Inversions 2L-7 and 3L-B were noted heterozygously in a very few specimens from other samples which were not scored systematically for autosomal polymorphisms.

- 2) Larva homozygous for inversion 2L-A..... *S. sanctipauli* 3)
- Inversion 2L-A absent from larva..... *S. soubrense* 4)
- 3) Inversion IS-21 Y-linked in population
..... *S. sanctipauli* Djodji form
Inversion IS-21 not Y-linked in population
..... *S. sanctipauli* typical form
- 4) Inversion 2S-6b absent from larva
..... *S. soubrense* typical form,
Inversion 2S-6c present in larva
..... *S. soubrense* Beffa form

Discussion

The new cytotype seems to be largely confined to the Asukaw-kaw and Dayi river systems in the mountainous forest on the Ghana/Togo border (see Table 1). Within Togo and Benin, to the north and east, *S. soubrense* Beffa form appears to be the sole representative of the *S. sanctipauli* subcomplex with the exception of a few samples of the Djodji form identified from the rivers Kpaza and Niankpe in October 1987. To the west of the Volta lake *S. sanctipauli* typical form and *S. soubrense* are found (Meredith et al. 1983, Post 1986, and Fiasorgbor, Weber, Post and Surtees, unpublished data).

In view of the absence of any sympatric samples or unique fixed inversions, there is no evidence for Djodji form being a species distinct from *S. sanctipauli* elsewhere. However, the sex-linkage of IS-21 indicates that Djodji populations are, by definition, genetically differentiated from other *S. sanctipauli* populations, and therefore it seems that Djodji form should be considered to be a geographic race within *S. sanctipauli*.

In a preliminary description this inversion was mistaken for a new inversion IS- α (Surtees 1986), because IS-21 is not just a simple reversal of the included bands, but is also associated with a consistent additional puff just outside the inversion and proximal to it. This gives IS-21 the superficial appearance of being longer in Djodji form.

The taxonomic significance of sex-linked inversions in the Simuliidae has been discussed by Post (1982), and within the *S. damnosum* complex sex-linked inversions have been considered important in the cytogenetic description of several forms and species, such as *S. soubrense* s.n. Beffa form (Meredith et al. 1983), *S. yahense* (Vajime and Dunbar 1975), and

Turiani form (Dunbar and Vajime 1981). In any case the importance of Djodji form, as with other forms described within the *S. damnosum* complex, lies not in its taxonomic level but rather in its possible epidemiological importance which is discussed by Garms and Cheke (1985) and Cheke and Denke (1988).

Acknowledgements

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