

Self-potential ambient noise and spectral relationship with urbanization, seismicity, and strain rate revealed via the Taiwan Geoelectric Monitoring Network

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Key Points:

- Self-potential (SP) signals show white noises or spectral ripples in areas of high industrialization and areas close to train stations.
- The SP spectral exponent seems to positively relate to the seismic b value and dilation rate and negatively to the urbanization level.
- The SP spectral exponent shows frequency-dependent correlations with the seismic b value, dilation rate, and urbanization level.

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Abstract

Geoelectric self-potential (SP) signals are sensitive to natural and anthropogenic factors. The SP spectral characteristics under the different factors in Taiwan were investigated and the SP spectral scalings were correlated with urbanization level, seismicity, and crustal deformation. The ambient SP noise models were first established by estimating the probability density functions of the spectrograms at each frequency. The effects of the natural and anthropogenic factors on the SP signals are understood by comparing the SP noise models under various conditions, such as precipitation, urbanization, and electric trains. Results show that the SP signals in areas of high industrialization and human activity and areas close to train stations behave as white noises and exhibit a distinct spectral ripple at frequencies around 1 Hz. On the other hand, the SP spectral power-law parameters, Gutenberg-Richter (GR) b values, and dilation strain rates were estimated by using the SP, earthquake catalog, and GPS data, respectively, during 2012–2017. By investigating the correlations of the SP spectral parameters with the GR b value, dilation strain rates, and urbanization level, the SP optimal frequency band is found between 0.006 and 1 Hz due to the high correlation between the SP and seismicity data and between the SP and dilation data and the low correlation between the SP and urbanization data. Hence, this study may help the filtering and screening of the SP data and facilitate the understanding of the mechano-electric behavior in the crust.

Plain Language Summary

Self-potential is a naturally electric potential difference beneath the Earth's surface, measured between any two points on the ground or in boreholes. The self-potential signals are sensitive to natural factors, such as ionospheric disturbance, solar wind, and tidal forces, and to artificial factors, such as electric trains, factories, and power pipelines. Hence, the investigation of the background noises is critical. The self-potential noise models were

established by estimating the probability density functions of the daily power spectral densities. We compared the noise models under several conditions, such as urbanization, electric trains, and rainfalls. We found that the noise models behave as white noises in areas of high industrialization and human activity. Furthermore, we also estimated the self-potential power-law parameters, seismic b values, and dilation strain rates. We then studied correlations of the self-potential power-law parameters with the urbanization levels, b values, and dilation strain rates. We found out the frequency-dependent correlations of the spectral scalings with the other parameters and determined that the self-potential signals with a high signal-to-noise ratio in the frequency band of 0.006–1 Hz. Understanding such correlations may understand the feature of the mechano-electric behavior in the crust and facilitate the development of a seismic-electric theory.

Keywords:

Self-potential, ambient noise model, urbanization level, Gutenberg-Richter b value, dilation strain rate.

1 Introduction

Self-potential (SP) is a naturally electric potential difference beneath the Earth's surface, measured between any two points on the ground or in boreholes. The SP signals occur in nature when electrolytes with different concentrations of ions come into galvanic contact with each other or where groundwater flows go through a capillary or a porous medium due to a hydraulic gradient [Revil & Jardani, 2013]. In recent decades, the SP method has been becoming an important component in geophysical surveys on the areas of groundwater flow [Jardani et al., 2006; Mao et al., 2015; Sill, 1983], dam seepage monitoring [Al-Saigh et al., 1994; Ikard et al., 2014; Panthulu et al., 2001], volcanic monitoring [Aizawa et al., 2005; Corwin & Hoover, 1979; Kuwano et al., 2015], and earthquake precursor [Chen & Chen, 2016; Chen et al., 2017; Corwin & Morrison, 1977; Ramírez-Rojas et al., 2004]. However, the SP signals are sensitive to anthropogenic factors, such as electric trains, factories, and power pipelines [Fraser-Smith & Coates, 1978; Harada et al., 2004; Ishikawa et al., 2007; Oettinger et al., 2001; Saito et al., 2011; Takahashi et al., 2007; Viljanen & Pirjola, 1994], and to natural factors, such as solar-terrestrial interactions, groundwater flows, and ocean tides [Fujii et al., 1995; Huang & Liu, 2006; Jardani et al., 2006; Lanzerotti et al., 2000]. Hence, understanding the SP background noises is critical and helpful to improve the data quality.

The SP signals clearly exhibit daily variations, which are affected by tidal forces and human activities [Huang & Liu, 2006; Szarka, 1988; Uyeda et al., 2002]. Hence, it is important to first understand the daily background noises, called ambient noises hereafter. However, it is extremely challenging to quantify the degree of human activities in certain areas, particularly for the degree of electricity usage, which is directly associated with the SP signals. Moreover, the SP signals are reported to behave as $1/f^\beta$ noises [Balasco et al., 2002; Ramírez-Rojas et al., 2011]. The Earth's crust is regarded as a self-organized critical system [Rundle et al., 2000 and references cited therein], exhibiting the fractal behaviors in

mechanical and electromagnetic ways [Chen et al., 2006; Potirakis et al., 2017; Sornette & Sornette, 1990]. Furthermore, crustal deformations are also suggested to affect the fractal behaviors [Öncel & Wilson, 2004]. Hence, it is attracting to study the relationships among the strain rates and scaling factors of the SP and seismicity data.

In this paper, we took Taiwan as an example and investigated correlations among the datasets of the SP, seismicity, strain rates, and urbanization. At present, the anthropogenic factor was provided with a detailed spatial distribution in Taiwan only for the urbanization data [Huang et al., 2018], which may relate to human activities regarding electricity usage. Taiwan is located at the collision boundary between the Eurasian plate and the Philippine Sea plate at a convergence rate of 82 mm/yr (Yu et al., 1997); hence, Taiwan has complicated geological features, active orogeny, and frequent earthquakes and provides multiple datasets concerning the SP, earthquake catalog, GPS (Global Positioning System), and urbanization data.

Since 2012, the Geoelectric Monitoring System (GEMS) has been established in Taiwan. It includes 20 SP stations about 50 km apart (Figure 1). One of the main purposes of the GEMS network is studying correlations between SP signals and crustal activities [Chen & Chen, 2016; Chen et al., 2017; Jiang et al., 2019; Telesca et al., 2014]. For instance, the SP anomalies were found to be associated with $M_L \geq 5$ earthquakes by using skewness and kurtosis and were attempted to create an earthquake forecasting algorithm [Chen & Chen, 2016; Chen et al., 2017]. Moreover, associations between the time dynamics of the SP signals and crustal deformation processes were investigated by using the Fisher-Shannon method [Telesca et al., 2014]. Obvious relations were found between the information properties of the SP signals and the differential strain intensity along the Taiwan orogen. However, due to a lack of ambient noise studies, it still remains contentious whether the GEMS network in Taiwan is producing meaningful SP data or this database is sufficient to study anomalous internally sourced fields.

In this study, we have characterized the SP $1/f^\beta$ noises by investigating the continuous data for 6 years from the GEMS network. There are two main objectives: (1) to establish the ambient SP noise models for each station by estimating the probability density functions of the spectrograms at each frequency [Berger et al., 2004; Marzorati & Bindi, 2006; McNamara & Buland, 2004]; and (2) to investigate the correlations of the SP spectral power-law models with the urbanization level, seismicity, and dilation strain rate. We then discussed the behaviors of the SP noise models under several conditions, such as urbanization, electric trains, precipitation. Also, we investigated the frequency dependence of the correlations of the SP spectral power-law parameters with the urbanization level, seismicity, and dilation strain rate. To our knowledge, despite a few studies concerning long-term electromagnetic data [Han et al., 2016; Huang, 2011; Xu et al., 2013], no research has focused on the ambient noises and spectral scaling correlations of SP data from a network with densely distributed, similarly equipped, and long-term recording stations. Consequently, this study may contribute to enrich the understanding of mechano-electric features in a complex environment like Taiwan.

2 Data

2.1 Geoelectric monitoring system and self-potential data

Since 2012, the GEMS network has been progressively installed in Taiwan and has been working by 20 SP stations uniformly distributed with a spacing of about 50 km (Figure 1). The SP data at each station are continuously recorded at a sampling rate of 15 samples per second and with a GPS time correction. The SP signals are the passive potential differences between two dipoles oriented in the North-South (NS) and East-West (EW) directions. The non-polarized electrodes (Pb–PbCl₂) of each dipole were buried at a depth of 1 m to avoid ground disturbances (e.g., surface temperature, precipitation, and human activity) and connected with the local telephone cable to each recording office. Table 1 lists the name,

location, and start date of each station. The information on the GEMS network can also be referred to a study [Chen & Chen, 2016].

In this paper, we analyzed the SP data for all components and stations measured from the onset time until 31 December 2017. Based on the 6-year data, we developed robust noise models for each SP station. To improve our understanding of the geoelectric features, we analyzed the daytime (09:00-14:00 Local Time, denoted as DT) and the nighttime (00:00-05:00 Local Time, denoted as NT) data, and compared them with the whole day (00:00-24:00 Local Time, denoted as WD) data. In contrast to the DT data, the effects of electric trains and most human activities can be well avoided for the NT data.

2.2 Urbanization data

The SP signals are sensitive to electric trains, factories, power pipelines, and other human activities concerning electricity usage [Fraser-Smith & Coates, 1978; Harada et al., 2004; Ishikawa et al., 2007; Oettinger et al., 2001; Saito et al., 2011; Takahashi et al., 2007]. However, it is exceptionally challenging to quantify those above-mentioned anthropogenic factors in certain regions. In this study, we employed the urbanization level estimated by Huang et al. [2018], who provided detailed spatial information on the degree of urbanization in Taiwan.

Urbanization is the process of the population moving from rural areas to urban areas, which is mainly proportional to the degree of human activities. Huang et al. [2018] studied the degree of urbanization in Taiwan townships and used data for the 2010 population, housing census, and so on. Four indicators in their study were introduced including “number of residents,” “the percentage of people working in secondary industries,” “the percentage of people working in tertiary industries,” and “population density.” Finally, the urbanization level was divided into five levels. Level 1 is the lowest, which had the lowest population size and density but the highest primary industry percentage. Level 5 is the highest, which had the

highest population size, population density, and the highest tertiary industry percentage. Detailed analyses and categories of the urbanization levels can be referred to the study [Huang et al., 2018]. The spatial distribution of the urbanization levels is mapped in Figure 1. The urbanization level of each SP station is also summarized in Table 1.

However, the urbanization level here is neither time-dependent during 2012–2017 nor directly related to the degree of electricity usage. In this study, we made two assumptions for comparisons between the urbanization and SP data. First, according to the Taiwan government reports (<https://www.stat.gov.tw/sitemap.asp>), the social change and economic growth have been becoming slow for the past decade; hence, we considered that the urbanization data for 2010 can be compared with the average parameters estimated from the datasets (e.g., SP, seismicity, and GPS) during 2012–2017. Second, we assumed that the urbanization level is proportional to the degree of the factors regarding electric trains, factories, and human activities regarding electricity usage. Therefore, the relationship between the electricity-related factors and SP data can be understood to some extent by studying the relationship between the urbanization and SP data.

2.3 Earthquake data

From 1973 to 1992, the Taiwan Telemetered Seismographic Network (TTSN) [Wang, 1989] was operated by the Institute of Earth Sciences, Academia Sinica. The TTSN consists of 24 stations, each equipped with a vertical high-gain and analog velocity seismometer. In 1992, the TTSN was merged into the Taiwan Seismic Network operated by the Central Weather Bureau (CWB) to form the CWB Seismic Network (CWBSN) [Shin, 1992; Shin et al., 2013]. The CWBSN is composed of 72 stations, each equipped with three-component digital velocity seismometers [Wu et al., 2008] and provided high-quality digital earthquake data that is recorded in both high- and low-gain formats. Since 2010, the CWBSN has again

been upgraded: the seismograms are currently digitized at 24 bits, and the number of stations has increased with more than 100 now widely installed in Taiwan [Chang et al., 2012; Hsiao et al., 2011]. The types of seismic instruments include short-period seismographs, accelerometers, and broadband seismometers. Therefore, the source parameters of an earthquake can be estimated more accurately and the completeness of magnitude (M_c) in this network can reach down to 1.5 [Chen et al., 2013].

That Earth's crust was observed to exhibit fractal behaviors in mechanical and electromagnetic ways [Chen et al., 2006; Potirakis et al., 2017]. Hence, we paid attention to study the relationship between the scaling factors of the seismicity and SP data. In this study, we employed the Taiwanese earthquake catalog during 2012–2017 with $M_L \geq 0$ in longitudes 119.5–122.5°E and latitudes 21.5–25.5°N. We then estimated the parameters of the Gutenberg-Richter (GR) law (Gutenberg & Richter, 1944; Rabinovitch et al., 2001), such as b and M_c values, by using the catalog selected within a 50-km radius for each SP station.

2.4 Global positioning system and data acquisition and processing

Crustal deformations affect the fractal behaviors in the crustal systems [Öncel & Wilson, 2004]. Hence, it is necessary to study the influence of deformation rates on the SP and GR scaling factors. In this study, we estimated dilation strain rates by using GPS data. Until 2019, the GPSLAB (<http://gps.earth.sinica.edu.tw>) has been managed more than 500 GPS stations operating in Taiwan. The GPS data were employed to perform numerous studies regarding crustal deformations and plate motions in Taiwan [Takahashi et al., 2019; Yu et al., 1997, 1999; Yu & Kuo, 2001].

Here, we analyzed the GPS data during 2012–2017 as the same observed period as the SP and earthquake data. The GPS data were downloaded from the GPSLAB and processed with the Jet Propulsion Laboratory's (JPL's) GIPSY-OASIS software [Bertiger et al., 2010; Zumberge

et al., 1997]. This software corrects the effects of the JPL's final orbit and clock information, the antenna phase center calibration, the first (and higher) order ionospheric bias, the tropospheric delays, the ocean loading, and so forth. These positions were computed under a new International GNSS Service (i.e., IGS08) reference frame [Rebischung et al., 2012]. Then, we estimated velocities and velocity uncertainties from the GPS coordinate time series with the Median Interannual Difference Adjusted for Skewness (MIDAS) algorithm [Blewitt et al., 2016]. This approach computes the median of slopes between all data pairs separated by one year, which is a variation of the Theil-Sen nonparametric median trend estimator [Sen, 1968; Theil, 1992]. The difference of pairs separated by one year mitigates the effect of annual periodicity rather than other transient signals such as different periodicities and co-seismic and post-seismic deformations. This technique has been shown to be more robust than the traditional least square with seasonal fitting [Blewitt et al., 2016] against common error sources such as step discontinuities, outliers, and seasonality.

3 Methodology

The SP signals are affected by daily variations, such as solar-terrestrial interactions, electric trains, and human activities regarding electricity usage. In order to understand the influences of natural or artificial sources on the SP signals, it is important to investigate the ambient SP noises, which is the daily background noises. In this study, we first established the ambient noise models for each SP station by computing the probability density functions (PDFs) on a set of the power spectral densities (PSDs) of selected SP signals. The methodology presented here was mainly referred to and modified from earlier studies [Berger et al., 2004; Marzorati & Bindi, 2006; McNamara & Buland, 2004], which applied the method for evaluating ambient earth noises. Hence, comparisons of the SP noise models can be carried out under several conditions, such as the urbanization level, precipitation, and electric trains (see Section 4).

On the other hand, the Earth's crust is a complex self-organizing system, which is

composed of multiple physical processes that interact between individuals to give rise to the overall state of the Earth as an entirety [Rundle et al., 2000; Turcotte, 1992]. Hence, the fractal behaviors in seismicity and electromagnetic data can be observed [Chen et al., 2006; Potirakis et al., 2017]. Also, the fractal behaviors are affected by crustal deformations [Öncel & Wilson, 2004]. Hence, in the second part, we analyzed and investigated the features of the SP $1/f^\beta$ noises, GR b values, and dilation strain rates. The processing steps and estimating methods are detailed as follows.

3.1 Self-potential data processing

The SP signals each day for every component and station were selected into one analyzed segment considering three kinds of time spans: (i) the whole day (00:00–24:00 Local Time) as the WD data, (ii) the daytime (09:00–14:00 Local Time) as the DT data, and (iii) the nighttime (00:00–05:00 Local Time) as the NT data. The effects of electric trains and most human activities can be well avoided for the NT data as a quiet reference. Figure 2a illustrates the original SP signals on 1 January 2016 for both components at CHCH. The SP amplitudes usually vary from tens to hundreds of mV/km with some extremely spiky signals. To better analyze the background noises, we thus discarded the data points with values greater than the 99.5th percentile and smaller than the 0.5th percentile in each segment. After such first data screening, if the percentage of missing data is less than 30% of one analyzed segment (at which the PSDs can still be well estimated by the Lomb-Scargle method [Munteanu et al., 2016], see Section 3.2), then the unevenly spaced time series are further processed. Next, we calculated the mean value and linear trend of each segment and subtracted them from that segment. The total number of the remaining segments for later estimating the PSDs using different time-spanned data at each station is listed in Table 1.

3.2 Power spectral density

The PSDs of the segments for different time-spanned data were estimated by the Welch method [Seatz et al., 2012; Welch, 1967] and the Lomb-Scargle method [Lomb, 1976; Munteanu et al., 2016; Scargle, 1982]. In each segment, we calculated the PSD in a moving window length of 10000 sec by using the Lomb-Scargle method with a shifting length of 1000 sec. Then, we computed a modified periodogram by averaging the estimated Lomb-Scargle PSDs in the segment. This averaged periodogram with overlapped time windows tends to reduce the variance for a single periodogram. This moving window is windowed by the Tukey window with a cosine taper of 25% width to decrease sidelobe leakage effects and to enhance the signal-to-noise ratio [Bloomfield, 2004]. The window length has 10000 sec so that the longest resolvable period is roughly 1000 sec (10^{-3} Hz) following an empirical relationship that the reliable longest period is about one-tenth of the calculating window [McNamara & Buland, 2004]. The frequencies are evenly spaced in a logarithmic scale with a step of 0.02 between 10^{-3} and $10^{0.86}$ ($\cong 7.24$) Hz. We then applied one-third octave smoothing to the PSD of each segment for reducing measurement noise [Hatziantoniu & Mourjopoulos, 2000; O'Haver, 2016; Tylka et al., 2017].

Let's take the SP time series shown in Fig. 2a as an example. Figure 2b demonstrates the PSD using the WD data for the NS component estimated by the above-mentioned method. We observed that the PSD between 0.001 and 0.1 Hz is approximately linear in a log-log scale. The linear fitting of the PSD was then estimated by using the least-squares regression (see Section 3.4). Furthermore, we estimated the daily PSDs using the different time-spanned data for all components and stations. Figure 3 illustrates the spectrogram at CHCH from its start date to 31 December 2017.

3.3 Probability density function

To investigate the ambient SP noises, we at each frequency calculated the PDF of the PSDs for the different time-spanned data. We grouped the PSD values into bins of 0.1

$\log_{10}(\text{mV}^2/\text{km}^2/\text{Hz})$ wide, ranging from -5 to 7, obtained histogram, and then divided by the total number of PSDs. Figures 4a and 4b demonstrate the PDFs for the WD data at frequencies of 0.003, 0.03, 0.3, and 3 Hz at CHCH. We observed that for both components the PDFs at lower frequencies, particularly when $f = 0.003$ Hz, are usually skewed to the high power densities. This implies that most of affecting factors on the SP signals seriously disturb the low-frequency signals. Figures 4c and 4d further demonstrate the probabilistic power spectral densities (PPSDs), which represent the PDFs for all frequencies in one plot. We also estimated and plotted the 5th percentile (P_5), mode (P_M), 95th percentile (P_{95}) at each frequency, which are referred to as the ambient SP noise models hereafter. In this case, the percentages around the $P_M(f)$ are greater than 12% for both components. Moreover, the distance of $P_{95}(f) - P_5(f)$ decreases with the frequency. This suggests that the SP signals at the low frequencies ($f < 0.1$ Hz) are more sensitive than those at the high frequencies ($f > 0.1$ Hz).

Figures 5a and 5b show the PPSD images using the DT data at CHCH for the NS and EW components, respectively. In contrast, Figs. 5c and 5d show the PPSD images using the NT data for the NS and EW components, respectively. The percentages on the $P_M(f)$ for the DT and NT data for both components are all greater than 12%. On the other hand, the distances of $P_{95}(f) - P_5(f)$ with $f < 0.1$ Hz for the NT data are smaller than those for the DT data. The day-night difference in this case logically suggests less interference of artificial noises during the nighttime. The distances of $P_{95}(f) - P_5(f)$ with $f > 10^{-1}$ Hz for the NT data for the EW component are wider than those for the NS component. This might be due to the induced telluric currents by the interference of the Earth's magnetic field or the telluric current circulation from land to sea [Simpson & Bahr, 2005; Zhang et al., 2019].

3.4 Scaling exponent of self-potential $1/f^\beta$ noise

The time dynamics of observational signals in the crustal system may exhibit correlation

structures and self-similarity in mechanical [Chen et al., 2006; Sarlis et al., 2015; Turcotte, 1999] or electromagnetic ways [Contoyiannis et al., 2004; Eftaxias et al., 2003; Potirakis et al., 2017; Rong et al., 2012; Varotsos et al., 2003], which generally indicate the inherent memory of the system generating those signals. Such correlation can be revealed by the power-law behavior of the PSD:

$$P(f) \propto 1/f^\beta. \quad (1)$$

The scaling exponent β is usually observed in the range $0 \leq \beta \leq 3$ and is generally estimated as the slope of the line fitting in the least-squares sense of Eq. (1) presented in log-log scales:

$$\log_{10}(P) = \alpha - \beta \cdot \log_{10}(f), \quad (2)$$

where α is a constant.

As a rule of thumb, a candidate power law should meaningfully exhibit a linear relationship in a log-log scale over approximately two orders of magnitude in both axes [Stumpf & Porter, 2012]. Hence, we selected a moving frequency window with 2 logarithmic orders with a shifting step of 0.2 logarithmic orders. For example, we fitted Eq. (2) to the PSDs between 10^{-3} and 10^{-1} Hz, between $10^{-2.8}$ and $10^{-0.8}$ Hz, and so on, until fitting between $10^{-1.2}$ and $10^{0.8}$ Hz. In this study, we attempted to understand whether there is a frequency-dependent relationship of the SP data with the urbanization data, GR b value, and dilation strain rate. Using the WD data by fitting Eq. (2) to the PSDs between 0.001 and 0.1 Hz, Fig. 6 illustrates at CHCH the daily variations of the estimated β values and the coefficient of determination (R^2) that quantifies how good the regression line fits the data. We observed that most of the β values for both components are in the range of 1–2 with $R^2 > 0.8$.

Further, we calculated the mean values and standard deviations of the daily estimated β and α values with $R^2 > 0.8$ for each SP station during 2012–2017. Table 2 shows the mean values of β and α (denoted as $\bar{\beta}$ and $\bar{\alpha}$) and the standard deviations of them (denoted as $\sigma_{\bar{\beta}}$ and $\sigma_{\bar{\alpha}}$) for all components and stations by using the different time-spanned data. In this case of fitting the PSDs between 0.001 and 0.1 Hz, Figs. 7a and 7b also illustrate the spatial

distributions of the $\bar{\beta}$ values using the NT data (denoted as $\bar{\beta}_{NT}$) for the NS and EW components, respectively. We observed that, for the different time-spanned data and both components, the $\bar{\beta}$ values for most of the SP stations range from 1 to 1.5 with $\sigma_{\bar{\beta}}$ from 0.2 to 0.4. This indicates that the ambient SP noises are often characterized by the flicker-noise and Brownian-noise processes [Guzmán-Vargas et al., 2009; Ramírez-Rojas et al., 2004]. On the other hand, the $\bar{\alpha}$ values for most of the SP stations range from -2 to 0 with $\sigma_{\bar{\alpha}}$ from 0.5 to 1.5. However, the scientific community was less interested in the discussion of the α values. In this paper, we attempted to discuss the related issue regarding the α values in the SP noises with the urbanization level, GR b value, and dilation strain rate.

3.5 Gutenberg-Richter b value

The frequency-magnitude distribution of seismicity is generally characterized by a power-law behavior [Gutenberg & Richter, 1944], also called the Gutenberg-Richter (GR) law:

$$\log_{10}[N(m \geq M_c)] = a - bm, \quad (3)$$

where $N(m \geq M_c)$ is the number of earthquakes with magnitude (m) larger than or equal to the completeness of magnitude (M_c), b is a scaling parameter, and a is a constant. Earthquake catalogs are shown to be incomplete at small magnitudes due to a detection threshold of seismic instruments and networks [Rydelek & Sacks, 1989]; hence, assessing the M_c value is critical. In this study, we estimated the b and M_c values by using the Clauset method [Clauset et al., 2009; Virkar & Clauset, 2014], which optimizes the lower data cutoff (i.e., completeness of magnitude in seismology) by minimizing the Kolmogorov-Smirnov distance between the empirical cumulative probability function (CDF) obtained from the data and the analytical CDF obtained via the GR-law parameters. The detailed application of the Clauset method can also be referred to a study [Li et al., 2015].

In this study, the b value was estimated by the earthquake catalog selected within a

50-km radius of each SP station, which is also the interstation distance. Table 1 shows the b value corresponding to each station. The errors in the b values were also calculated by using a bootstrap approach [Schorlemmer et al., 2003]. In this approach, we randomly resampled a dataset of the same number of events on the selected catalog for each station and used the Clauset method to estimate the b value of the resampled dataset. We repeated this procedure 1000 times and took the mean value and standard deviation to represent the confidence in b (denoted as \bar{b} and $\sigma_{\bar{b}}$). As Figure 7c shows the spatial distribution of the estimated \bar{b} values, which are close to 1 for most areas. However, the middle and western sides of Taiwan are characterized by smaller \bar{b} values (below 0.9) than the eastern side. Furthermore, the northern sides are featured by higher \bar{b} values (above 1.1); particularly, the \bar{b} value ($b = 1.74 \pm 0.29$) is maximum at SHRL located on the Tatun Volcano Group, in agreement that volcanic areas are usually featured by higher b values [Pu et al., 2014; Wiemer & McNutt, 1997; Wyss et al., 1997].

3.6 Dilation strain rate

The dilation strain rate (D) is the first invariant of the strain rate tensor ($\dot{\epsilon}_{ij}$), which equals the trace of $\dot{\epsilon}_{ij}$:

$$D = \dot{\epsilon}_{xx} + \dot{\epsilon}_{yy} + \dot{\epsilon}_{zz}. \quad (4)$$

Since the GPS does not provide reliable vertical velocities [Schmidt et al., 2003; Wing & Frank, 2011], we considered the areal dilation rate, which is the sum of the maximum and minimum principal strain rates ($\dot{\epsilon}_1$ and $\dot{\epsilon}_2$, respectively):

$$D = \dot{\epsilon}_1 + \dot{\epsilon}_2. \quad (5)$$

The positive and negative D values correspond, respectively, to extension or elongation and contraction or shortening. We derived strain rates from the GPS velocities based on the VISR (Velocity Interpolation to Strain Rates) method described by Shen et al. [2015]. In this method, we meshed the Taiwan area and set a grid size of 0.05×0.05 degree. The local strain

rate at each node of the grid was estimated from local geodetic measurements using Gaussian and Voronoi cell weightings to automatically reduce the contribution of far stations from the node within the entire network. A uniform strain rate field was assumed at each node of the grid and a weighted least squares inversion was performed over strain component solutions and their covariance for the unknowns of translation, rotation, and strain. Next, we estimated the derived quantities of the strain rates such as the principal strain rates and dilation strain rates.

We then selected the nodes of the grid within a 20-km radius of each SP station and averaged the dilation strain rates of the selected nodes. Table 1 lists the mean values and standard deviations of the dilation strain rates (denoted as \bar{D} and $\sigma_{\bar{D}}$). Figure 7d shows the spatial distribution of the \bar{D} values. The large positive \bar{D} values (above 0.2 $\mu\text{strain/yr}$) feature around TOCH on account of the back-arc rifting or extension in the Okinawa Trough [Liu et al., 2016; Sibuet et al., 1987]. On the other hand, the large negative \bar{D} values (below -1.6 $\mu\text{strain/yr}$) appear at YULI and RUEY due to the strong convergence of the Luzon arc [Mouthereau et al., 2009; S. Yu et al., 1999].

4 Results

4.1 Ambient self-potential noise models

In this section, we paid attention to the comparisons of the SP noise models (P_5 , P_M , and P_{95}) under different conditions, such as the urbanization level, electric trains (associated with the day-night difference), and precipitation.

To understand the influence of the urbanization level on the SP data, we compared the SP noise models for the WD data at KAOH with those at HUAL, as shown in Fig. 8. Station KAOH ($U = 4$) is situated in highly industrialized areas in Taiwan and close to the area of $U = 5$, while HUAL ($U = 1$) is located in desolate mountain areas as a silent spatial reference. We observed that the SP noise models for both components at KAOH are flatter than those at

HUAL, which are close to white noises. Moreover, distinct ripples appear in the noise models at KAOH, particularly at frequencies of 0.01 and 1 Hz. These results indicate the serious interference by human activities at KAOH, the highly industrialized area.

Second, to understand the effect of the electric trains on the SP data, we compared the SP noise models for the DT data with those for the NT data at HERM and HUAL, as shown in Fig. 9. Station HERM is close to a busy train station with a distance of 5 km, whereas there is no train effect at HUAL in remote mountains. Additionally, the DT data is spanned during the working hours of 09:00–14:00 Local Time; contrarily, the NT data is spanned during 00:00–05:00 Local Time, when the train effect can be totally avoided, as a quiet temporal reference. We observed that, at HERM and for both components, the SP noise models when $f > 10^{-0.5}$ Hz during the nighttime are $0.5 \log_{10}(\text{mV}^2/\text{km}^2/\text{Hz})$ smaller than those during the daytime. Exceptionally, the P_{95} models around $f \cong 1$ Hz feature a distinct ripple and have the same order for both DT and NT data. Moreover, the distances of $P_{95}(f) - P_5(f)$ with $f < 0.01$ Hz during the nighttime are smaller than those during the daytime. As the quiet reference at HUAL, the SP noise models are overall similar for both the DT and NT data. Only the P_{95} model for the NS component during the nighttime is smaller than that during the daytime.

Third, to understand the impact of precipitation on the SP data, we compared the SP noise models at the quiet station HUAL and using the WD data, which were re-estimated by the proposed methodology under the conditions of precipitation equal to 0 mm/day, greater than 50 mm/day, and greater than 100 mm/day. Figure 10a shows the daily time series of precipitation at HUAL, which displays the seasonal variations that the precipitation amounts are highly focused in summer times. Figures 7b and 7c show the ambient noise models with precipitation equal to 0 mm/day and greater than 50 mm/day for the NS and EW components, respectively. Here the SP noise models with precipitation equal to 0 mm/day are served as a dry reference. Figures 7d and 7e also show the SP noise models with precipitation equal to 0

mm/day and greater than 100 mm/day for the NS and EW components, respectively. We observed that the P_M models are almost the same for the three precipitation conditions. However, the P_5 model for the NS component with precipitation > 100 mm/day has larger PSDs for all frequencies than that with precipitation > 50 mm/day, which has also larger PSDs than that with precipitation of 0 mm/day. Moreover, the P_{95} model for the EW component with precipitation > 100 mm/day has larger PSD when $f > 10^{-0.7}$ Hz than that with precipitation > 50 mm/day, which has also larger PSDs than that with precipitation of 0 mm/day. Overall, whether the precipitation is considered, the ambient SP noise models show little difference. The reason may be that long dipole lengths are used (1–3 km) in the GEMS network. While the precipitation can cause a prominent flow direction in small scales and the strong effect of the precipitation on the SP signals in short dipole lengths such as tens of meters, the GEMS data look immune to rainfalls.

4.2 Correlations among self-potential power-law parameters, urbanization level, Gutenberg-Richter b value, and dilation strain rate

In this section, we focused on analyzing spatial correlations (symbolized as ρ hereafter) among the urbanization level (U), GR \bar{b} value, dilation strain rate (\bar{D}), and SP power-law parameters ($\bar{\alpha}$ and $\bar{\beta}$) using the different time-spanned data and for both components. By the way, we simply denoted the correlation between U and \bar{b} as $\rho_{U\bar{b}}$, for instance.

Figure 11 illustrates the correlation and corresponding p-value matrices when the $\bar{\alpha}$ and $\bar{\beta}$ values were estimated in the frequency band of 0.001–0.1 Hz. The correlation matrices for the NS and EW components are imaged as Figs. 11a and 11c, respectively; correspondingly, the p-value matrices for the NS and EW components are imaged as Figs. 11b and 11d, respectively. We observed that the absolutes of correlations with $|\rho| > 0.1$ have significance with $p \leq 0.1$. The correlations of $\rho_{\bar{\alpha}WD\bar{\alpha}DT}$, $\rho_{\bar{\alpha}WD\bar{\alpha}NT}$, and $\rho_{\bar{\alpha}DT\bar{\alpha}NT}$ are highly correlated, and so are those of $\rho_{\bar{\beta}WD\bar{\beta}DT}$, $\rho_{\bar{\beta}WD\bar{\beta}NT}$, and $\rho_{\bar{\beta}DT\bar{\beta}NT}$. This logically indicates that the SP

power-law models at each station have similar patterns whether they are estimated using the WD, DT, or NT data. No spatial correlation exists between U and \bar{D} ($\rho_{U\bar{D}} = 0.1$), which is reasonable that the crustal deformations are mainly caused by plate motions and mantle convection [Pysklywec & Beaumont, 2004; S. Yu et al., 1999]. However, a moderate correlation seems to exist between U and \bar{b} ($\rho_{U\bar{b}} = 0.39$) (see Section 5). Moreover, a moderate correlation might also exist between \bar{b} and \bar{D} ($\rho_{\bar{b}\bar{D}} = 0.41$) in agreement with the finding that seismicity is more common at positive strain rate values [Hauksson, 2011; Nakamura & Kinjo, 2018].

We now concentrated on the correlations of the SP power-law parameters with the urbanization level, GR \bar{b} value, and dilation strain rate. In this case of $\bar{\alpha}$ and $\bar{\beta}$ estimated between 0.001 and 0.1 Hz, we observed that the correlations between U and $\bar{\beta}$ for the different time-spanned data are featured by relatively high negative values (-0.65 to -0.47) for both components. This means that larger U values correspond to smaller $\bar{\beta}$ values, that is reminiscent of the case of KAOH (Fig. 8). We considered that the SP signals are seriously dominated by white noises when an SP station is located in highly urbanized areas. However, in this case, the correlations of $\bar{\alpha}$ and $\bar{\beta}$ with \bar{b} and \bar{D} almost do not exist ($|\rho| \leq 0.3$).

To further understand the influence of the urbanization level, GR \bar{b} value, and dilation strain rate on the frequency-dependent SP power-law parameters, we then analyzed the correlations of U , \bar{b} and \bar{D} with $\bar{\alpha}$ and $\bar{\beta}$ estimated in different frequency bands (see Section 3.4). Figure 12 shows the correlations of U , \bar{b} and \bar{D} with $\bar{\alpha}$ estimated in different frequency bands for both components using the different time-spanned data. We observed that, no matter what frequency bands are involved, most of them are not correlated ($|\rho| \leq 0.3$). Exceptionally, the correlations between U and $\bar{\alpha}_{DT}$ at some frequency bands can be featured by large negative values (below -3). This again implies that high human activities during the daytime indeed seriously affect the SP signals. On the other hand, Fig. 13 similarly shows the correlations of U , \bar{b} and \bar{D} with $\bar{\beta}$ estimated in different frequency bands. We

observed that, at certain frequency bands, moderate correlations of U , \bar{b} and \bar{D} with $\bar{\beta}$ can reach values of $|\rho| > 0.3$. This result indicates that the correlations are frequency-dependent. Compared those curves, the optimal frequency bands ($10^{-2.2}$ to 1 Hz) can be found due to higher positive correlations of $\rho_{\bar{b}\bar{\beta}}$ and $\rho_{\bar{D}\bar{\beta}}$ and less correlation of $\rho_{U\bar{\beta}}$.

5 Discussion and conclusions

The effects of anthropogenic factors (e.g., electric trains, factories, and power pipelines) are reported to disturb electromagnetic signals [Fraser-Smith & Coates, 1978; Ishikawa et al., 2007; Saito et al., 2011; Viljanen & Pirjola, 1994]. However, it is difficult to quantify such anthropogenic factors in certain areas. Here we choose the detailed spatial distribution of the urbanization level in Taiwan estimated by Huang et al. [2018] on behalf of the anthropogenic factors. In this paper, we demonstrated the comparisons of the ambient SP noise models under different artificial factors, such as the urbanization level and electric trains. As Fig. 8 shows, the comparison for the urbanization level suggests that the SP signals are characterized by white noise processes in areas of high industrialization or human activity. On the other hand, in Fig. 9, the day-night difference, which mostly responds to the train effect, between the SP noise models indicates that the SP signals with $f > 10^{-0.5}$ ($\cong 0.32$) Hz and $f < 0.01$ Hz are mainly affected by the train noises during the daytime. Hence, installing one SP station in the above-mentioned areas should be avoided. However, due to site limitations, the SP signals are sometimes inevitably disturbed by the anthropogenic factors. Logically, a further and critical step is to develop a noise reduction method or apply developed methods to distinguish noises from signals, such as remote reference [Larsen, 1997; Oettinger et al., 2001] and interstation transfer function [Harada et al., 2004; Takahashi et al., 2007]. The ambient SP noise models proposed here play an important role to examine whether noises are successfully extracted from signals after using noise reduction methods. For example, we would expect that the SP noise models become flicker-noise or

Brownian-noise patterns and the spectral ripples in the noise models disappear if one applied the interstation transfer function approach to the SP signals at KAOH. At present, this remains an open question and will be our future work.

The most important topic is to study the frequency dependence of the correlations of the SP power-law parameters (α and β) with the urbanization level (U), GR b value, and dilation strain rate (D). As Fig. 12 shows, the α values are not correlated with the U , b , and D values. However, the α represents the energy level of the analyzed signals; hence, we believed that the α values in the SP power-law models should be related to the geological settings or hydrological conditions, which can affect the electromagnetic signals [Befus et al., 2014; Rizzo et al., 2004]. On the other hand, the correlations of β with U , b , and D show the frequency dependence, and the optimal frequency band is found between $10^{-2.2}$ ($\cong 0.006$) and 1 Hz, at which there are relatively high correlations between β and b and between β and D and relatively low correlation between β and U . The Earth's crust, as a complex dynamic system, exhibits the fractal behaviors in the mechanical and electromagnetic ways in the same time [Rabinovitch et al., 2001]; hence, the fractals in the seismicity and SP data logically have a positive correlation, which reaches a maximum of 0.6 in this study. Moreover, the correlation between β and D reaches a maximum of 0.4. As Öncel & Wilson [2004] pointed out, the crustal deformations can affect the fractal behaviors in the crust. Therefore, there are positive correlations among β , b , and D in this study. However, it is still unsolved that the effects of the anthropogenic and crustal factors on the SP signals between 0.006 and 1 Hz are better performed.

In the investigation of the spatial correlation matrix, interestingly, the positive correlation between the urbanization level and GR b value seems to exist. The urbanization level is mainly related to the degree of population, which is related to the surface mass loading. Hence, the high urbanization level represents large surface mass loading. It has been reported that the surface mass loading may affect the seismicity, such as snow load [Heki,

2003] and reservoir mass [Mandal et al., 2005; Simpson, 1976]. Additionally, the coverage and detection of the seismic network should be considered. The seismometers in Taiwan are densely distributed in cities (high urbanization level) rather than in mountain areas (low urbanization level) [Shin et al., 2013]; hence, more small earthquakes in highly urbanized areas may be recorded and the GR b values are then affected to increase. However, there is no rigorous analysis and inference in this paper; hence, it is still arguable that earthquakes may be triggered by the population mass loading or caused due to the detection capability of the seismic network.

In conclusion, we have investigated the ambient SP noise models under different conditions, such as the urbanization level, electric trains, and precipitation. Understanding the patterns of the ambient SP noises is helpful to maintain the original SP stations or deploy new stations. As a typical example, the SP signals at KAOH are atrocious. Second, we have investigated the spatial correlation matrix and its frequency dependence of the SP power-law parameters with the urbanization level, GR b value, and dilation strain rate. The determination of the optimal frequency band allows us to filter and screen the SP signals and improve the quality of the SP analyses. Consequently, this paper may be conducive to understand the seismo-electric theory and mechano-electric features in the crust.

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References

- Aizawa, K., Yoshimura, R., Oshiman, N., Yamazaki, K., Uto, T., Ogawa, Y., et al. (2005). Hydrothermal system beneath Mt. Fuji volcano inferred from magnetotellurics and electric self-potential. *Earth and Planetary Science Letters*, 235(1), 343–355. <https://doi.org/10.1016/j.epsl.2005.03.023>
- Al-Saigh, N. H., Mohammed, Z. S., & Dahham, M. S. (1994). Detection of water leakage from dams by self-potential method. *Engineering Geology*, 37(2), 115–121. [https://doi.org/10.1016/0013-7952\(94\)90046-9](https://doi.org/10.1016/0013-7952(94)90046-9)
- Balasco, M., Lapenna, V., & Telesca, L. (2002). $1/f^\alpha$ Fluctuations in geoelectrical signals observed in a seismic area of Southern Italy. *Tectonophysics*, 347(4), 253–268. [https://doi.org/10.1016/S0040-1951\(02\)00062-8](https://doi.org/10.1016/S0040-1951(02)00062-8)
- Befus, K. M., Cardenas, M. B., Tait, D. R., & Erler, D. V. (2014). Geoelectrical signals of geologic and hydrologic processes in a fringing reef lagoon setting. *Journal of Hydrology*, 517, 508–520. <https://doi.org/10.1016/j.jhydrol.2014.05.070>
- Berger, J., Davis, P., & Ekström, G. (2004). Ambient Earth noise: A survey of the Global Seismographic Network. *Journal of Geophysical Research: Solid Earth*, 109(B11), B11307. <https://doi.org/10.1029/2004JB003408>
- Bertiger, W., Desai, S. D., Haines, B., Harvey, N., Moore, A. W., Owen, S., & Weiss, J. P. (2010). Single receiver phase ambiguity resolution with GPS data. *Journal of*

Geodesy, 84(5), 327–337. <https://doi.org/10.1007/s00190-010-0371-9>

Blewitt, G., Kreemer, C., Hammond, W. C., & Gazeaux, J. (2016). MIDAS robust trend estimator for accurate GPS station velocities without step detection. *Journal of Geophysical Research: Solid Earth*, 121(3), 2054–2068. <https://doi.org/10.1002/2015JB012552>

Bloomfield, P. (2004). *Fourier Analysis of Time Series: An Introduction*. John Wiley & Sons.

Chang, C.-H., Wu, Y.-M., Chen, D.-Y., Shin, T.-C., Chin, T.-L., & Chang, W.-Y. (2012). An Examination of Telemetry Delay in the Central Weather Bureau Seismic Network. *Terrestrial, Atmospheric and Oceanic Sciences*, 23(3), 261. [https://doi.org/10.3319/TAO.2011.11.29.01\(T\)](https://doi.org/10.3319/TAO.2011.11.29.01(T))

Chen, C.-C., Wang, W.-C., Chang, Y.-F., Wu, Y.-M., & Lee, Y.-H. (2006). A correlation between the b-value and the fractal dimension from the aftershock sequence of the 1999 Chi-Chi, Taiwan, earthquake. *Geophysical Journal International*, 167(3), 1215–1219. <https://doi.org/10.1111/j.1365-246X.2006.03230.x>

Chen, C.-H., Wang, J.-P., Wu, Y.-M., Chan, C.-H., & Chang, C.-H. (2013). A study of earthquake inter-occurrence times distribution models in Taiwan. *Natural Hazards*, 69(3), 1335–1350. <https://doi.org/10.1007/s11069-012-0496-7>

Chen, H.-J., & Chen, C.-C. (2016). Testing the correlations between anomalies of statistical indexes of the geoelectric system and earthquakes. *Natural Hazards*, 84(2), 877–895.

<https://doi.org/10.1007/s11069-016-2460-4>

Chen, H.-J., Chen, C.-C., Ouillon, G., & Sornette, D. (2017). Using skewness and kurtosis of geoelectric fields to forecast the 2016/2/6, ML6.6 Meinong, Taiwan Earthquake. *Terrestrial, Atmospheric and Oceanic Sciences*, 28(5), 745–761.

<https://doi.org/10.3319/TAO.2016.11.01.01>

Clauset, A., Shalizi, C., & Newman, M. (2009). Power-Law Distributions in Empirical Data. *SIAM Review*, 51(4), 661–703. <https://doi.org/10.1137/070710111>

Contoyiannis, Y. F., Diakonos, F. K., Kapiris, P. G., Peratzakis, A. S., & Eftaxias, K. A. (2004). Intermittent dynamics of critical pre-seismic electromagnetic fluctuations.

Physics and Chemistry of the Earth, Parts A/B/C, 29(4), 397–408.

<https://doi.org/10.1016/j.pce.2003.11.012>

Corwin, R. F., & Hoover, D. B. (1979). The self-potential method in geothermal exploration. *Geophysics*, 44(2), 226–245. <https://doi.org/10.1190/1.1440964>

Corwin, R. F., & Morrison, H. F. (1977). Self-potential variations preceding earthquakes in central California. *Geophysical Research Letters*, 4(4), 171–174.

<https://doi.org/10.1029/GL004i004p00171>

Eftaxias, K., Kapiris, P., Polygiannakis, J., Peratzakis, A., Kopanas, J., Antonopoulos, G., & Rigas, D. (2003). Experience of short term earthquake precursors with VLF–VHF electromagnetic emissions. *Natural Hazards and Earth System Sciences*, 3(3/4),

217–228. <https://doi.org/10.5194/nhess-3-217-2003>

Fraser-Smith, A. C., & Coates, D. B. (1978). Large-amplitude ULF electromagnetic fields from BART. *Radio Science*, 13(4), 661–668.

<https://doi.org/10.1029/RS013i004p00661>

Fujii, I., Lanzerotti, L. J., Utada, H., Kinoshita, H., Kasahara, J., Medford, L. V., & MacLennan, C. G. (1995). Geoelectric power spectra over oceanic distances.

Geophysical Research Letters, 22(4), 421–424.

[https://doi.org/10.1029/94GL03282@10.1002/\(ISSN\)1542-7390.GIC15](https://doi.org/10.1029/94GL03282@10.1002/(ISSN)1542-7390.GIC15)

Gutenberg, B., & Richter, C. F. (1944). Frequency of earthquakes in California. *Bulletin of the Seismological Society of America*, 34(4), 185–188.

Guzmán-Vargas, L., Ramírez-Rojas, A., Hernández-Pérez, R., & Angulo-Brown, F. (2009).

Correlations and variability in electrical signals related to earthquake activity. *Physica A: Statistical Mechanics and Its Applications*, 388(19), 4218–4228.

<https://doi.org/10.1016/j.physa.2009.06.019>

Han, P., Hattori, K., Huang, Q., Hirooka, S., & Yoshino, C. (2016). Spatiotemporal characteristics of the geomagnetic diurnal variation anomalies prior to the 2011

Tohoku earthquake (Mw 9.0) and the possible coupling of multiple pre-earthquake phenomena. *Journal of Asian Earth Sciences*, 129, 13–21.

<https://doi.org/10.1016/j.jseaes.2016.07.011>

Harada, M., Hattori, K., & Isezaki, N. (2004). Transfer function approach to signal discrimination of ULF geomagnetic data. *Physics and Chemistry of the Earth, Parts A/B/C*, 29(4), 409–417. <https://doi.org/10.1016/j.pce.2004.03.002>

Hatziantoniou, P. D., & Mourjopoulos, J. N. (2000). Generalized Fractional-Octave Smoothing of Audio and Acoustic Responses. *Journal of the Audio Engineering Society*, 48(4), 259–280.

Hauksson, E. (2011). Crustal geophysics and seismicity in southern California. *Geophysical Journal International*, 186(1), 82–98. <https://doi.org/10.1111/j.1365-246X.2011.05042.x>

Heki, K. (2003). Snow load and seasonal variation of earthquake occurrence in Japan. *Earth and Planetary Science Letters*, 207(1–4), 159–164. [https://doi.org/10.1016/S0012-821X\(02\)01148-2](https://doi.org/10.1016/S0012-821X(02)01148-2)

Hsiao, N.-C., Wu, Y.-M., Zhao, L., Chen, D.-Y., Huang, W.-T., Kuo, K.-H., et al. (2011). A new prototype system for earthquake early warning in Taiwan. *Soil Dynamics and Earthquake Engineering*, 31(2), 201–208. <https://doi.org/10.1016/j.soildyn.2010.01.008>

Huang, N.-C., Kung, S.-F., & Hu, S. C. (2018). The Relationship between Urbanization, the Built Environment, and Physical Activity among Older Adults in Taiwan. *International Journal of Environmental Research and Public Health*, 15(5), 20 pp.

<https://doi.org/10.3390/ijerph15050836>

Huang, Q. (2011). Retrospective investigation of geophysical data possibly associated with the Ms8.0 Wenchuan earthquake in Sichuan, China. *Journal of Asian Earth Sciences*, 41(4), 421–427. <https://doi.org/10.1016/j.jseaes.2010.05.014>

Huang, Q. H., & Liu, T. (2006). Earthquakes and tide response of geoelectric potential field at the Nijima station. *Chinese Journal of Geophysics*, 49(6), 1745–1754.

Ikard, S. J., Revil, A., Schmutz, M., Karaoulis, M., Jardani, A., & Mooney, M. (2014). Characterization of Focused Seepage Through an Earthfill Dam Using Geoelectrical Methods. *Groundwater*, 52(6), 952–965. <https://doi.org/10.1111/gwat.12151>

Ishikawa, H., Hattori, K., Takahashi, I., Noda, Y., Nagao, T., & Isezaki, N. (2007). Effect of Noise from DC-Driven Trains to Geoelectrical Potential Difference and its Reduction in Hakuba Area, Japan. *IEEJ Transactions on Fundamentals and Materials*, 127(1), 41–47. <https://doi.org/10.1541/ieejfms.127.41>

Jardani, A., Dupont, J. P., & Revil, A. (2006). Self-potential signals associated with preferential groundwater flow pathways in sinkholes. *Journal of Geophysical Research: Solid Earth*, 111(B9). <https://doi.org/10.1029/2005JB004231>

Jiang, F., Chen, X., Chen, C.-C., & Chen, H.-J. (2019). Relationship Between Seismic Electric Signals and Tectonics Derived from Dense Geoelectric Observations in Taiwan. *Pure and Applied Geophysics*. <https://doi.org/10.1007/s00024-018-2077-5>

Kuwano, O., Yoshida, S., Nakatani, M., & Uyeshima, M. (2015). Origin of transient self-potential signals associated with very long period seismic pulses observed during the 2000 activity of Miyakejima volcano. *Journal of Geophysical Research: Solid Earth*, *120*, 3544–3565.

[https://doi.org/10.1002/2014JB011740@10.1002/\(ISSN\)2169-9356.SSMCV1](https://doi.org/10.1002/2014JB011740@10.1002/(ISSN)2169-9356.SSMCV1)

Lanzerotti, L. J., Sayres, D. S., Medford, L. V., MacLennan, C. G., Lepping, R. P., & Szabo, A. (2000). Response of large-scale geoelectric fields to identified interplanetary disturbances and the equatorial ring current. *Advances in Space Research*, *26*(1), 21–26. [https://doi.org/10.1016/S0273-1177\(99\)01021-2](https://doi.org/10.1016/S0273-1177(99)01021-2)

Larsen, J. C. (1997). Noise Reduction in Electromagnetic Time Series to Improve Detection of Seismic-Induced Signals. *Journal of Geomagnetism and Geoelectricity*, *49*(11–12), 1257–1265. <https://doi.org/10.5636/jgg.49.1257>

Li, H.-C., Chang, C.-H., & Chen, C.-C. (2015). Quantitative Analysis of Seismicity Before Large Taiwanese Earthquakes Using the G-R Law. *Pure and Applied Geophysics*, *173*(1), 153–164. <https://doi.org/10.1007/s00024-014-1030-5>

Liu, B., Li, S.-Z., Suo, Y.-H., Li, G.-X., Dai, L.-M., Somerville, I. D., et al. (2016). The geological nature and geodynamics of the Okinawa Trough, Western Pacific. *Geological Journal*, *51*(S1), 416–428. <https://doi.org/10.1002/gj.2774>

Lomb, N. R. (1976). Least-squares frequency analysis of unequally spaced data. *Astrophysics*

and Space Science, 39(2), 447–462. <https://doi.org/10.1007/BF00648343>

Mandal, P., Mabawonku, A. O., & Dimri, V. P. (2005). Self-organized Fractal Seismicity of Reservoir Triggered Earthquakes in the Koyna-Warna Seismic Zone, Western India. *Pure and Applied Geophysics*, 162(1), 73–90. <https://doi.org/10.1007/s00024-004-2580-8>

Mao, D., Revil, A., Hort, R. D., Munakata-Marr, J., Atekwana, E. A., & Kulesa, B. (2015). Resistivity and self-potential tomography applied to groundwater remediation and contaminant plumes: Sandbox and field experiments. *Journal of Hydrology*, 530, 1–14. <https://doi.org/10.1016/j.jhydrol.2015.09.031>

Marzorati, S., & Bindi, D. (2006). Ambient noise levels in north central Italy. *Geochemistry, Geophysics, Geosystems*, 7(9). <https://doi.org/10.1029/2006GC001256>

McNamara, D. E., & Buland, R. P. (2004). Ambient Noise Levels in the Continental United States. *Bulletin of the Seismological Society of America*, 94(4), 1517–1527. <https://doi.org/10.1785/012003001>

Mouthereau, F., Fillon, C., & Ma, K.-F. (2009). Distribution of strain rates in the Taiwan orogenic wedge. *Earth and Planetary Science Letters*, 284(3–4), 361–385. <https://doi.org/10.1016/j.epsl.2009.05.005>

Munteanu, C., Negrea, C., Echim, M., & Mursula, K. (2016). Effect of data gaps: comparison of different spectral analysis methods. *Annales Geophysicae*, 34(4), 437–449.

<https://doi.org/10.5194/angeo-34-437-2016>

Nakamura, M., & Kinjo, A. (2018). Activated seismicity by strain rate change in the Yaeyama region, south Ryukyu. *Earth, Planets and Space*, 70(1), 154.

<https://doi.org/10.1186/s40623-018-0929-y>

Oettinger, G., Haak, V., & Larsen, J. C. (2001). Noise reduction in magnetotelluric time-series with a new signal–noise separation method and its application to a field experiment in the Saxonian Granulite Massif. *Geophysical Journal International*, 146(3), 659–669. <https://doi.org/10.1046/j.1365-246X.2001.00473.x>

O’Haver, T. (2016). *A Pragmatic Introduction to Signal Processing*. Lulu Enterprises Incorporated.

Öncel, A. O., & Wilson, T. (2004). Correlation of seismotectonic variables and GPS strain measurements in western Turkey. *Journal of Geophysical Research: Solid Earth*, 109(B11), B11306. <https://doi.org/10.1029/2004JB003101>

Panthulu, T. V., Krishnaiah, C., & Shirke, J. M. (2001). Detection of seepage paths in earth dams using self-potential and electrical resistivity methods. *Engineering Geology*, 59(3), 281–295. [https://doi.org/10.1016/S0013-7952\(00\)00082-X](https://doi.org/10.1016/S0013-7952(00)00082-X)

Potirakis, S. M., Hayakawa, M., & Schekotov, A. (2017). Fractal analysis of the ground-recorded ULF magnetic fields prior to the 11 March 2011 Tohoku earthquake (M_w = 9): discriminating possible earthquake precursors from space-sourced

disturbances. *Natural Hazards*, 85(1), 59–86.

<https://doi.org/10.1007/s11069-016-2558-8>

Pu, H.-C., Lin, C.-H., Huang, Y.-C., Chang, L.-C., Lee, H.-F., Leu, P.-L., et al. (2014). The Volcanic Earthquake Swarm of October 20, 2009 in the Tatun Area of Northern Taiwan. *Terrestrial, Atmospheric and Oceanic Sciences*, 25(5), 625.

[https://doi.org/10.3319/TAO.2014.04.11.02\(T\)](https://doi.org/10.3319/TAO.2014.04.11.02(T))

Pysklywec, R. N., & Beaumont, C. (2004). Intraplate tectonics: feedback between radioactive thermal weakening and crustal deformation driven by mantle lithosphere instabilities.

Earth and Planetary Science Letters, 221(1), 275–292.

[https://doi.org/10.1016/S0012-821X\(04\)00098-6](https://doi.org/10.1016/S0012-821X(04)00098-6)

Rabinovitch, A., Frid, V., & Bahat, D. (2001). Gutenberg-Richter-type relation for laboratory fracture-induced electromagnetic radiation. *Physical Review E*, 65(1), 011401.

<https://doi.org/10.1103/PhysRevE.65.011401>

Ramírez-Rojas, A., Telesca, L., & Angulo-Brown, F. (2011). Entropy of geoelectrical time series in the natural time domain. *Natural Hazards and Earth System Sciences*, 11(1),

219–225. <https://doi.org/10.5194/nhess-11-219-2011>

Ramírez-Rojas, A., Pavía-Miller, C. G., & Angulo-Brown, F. (2004). Statistical behavior of the spectral exponent and the correlation time of electric self-potential time series associated to the Ms=7.4 September 14, 1995 earthquake in Mexico. *Physics and*

Chemistry of the Earth, Parts A/B/C, 29(4), 305–312.

<https://doi.org/10.1016/j.pce.2003.10.001>

Rebischung, P., Griffiths, J., Ray, J., Schmid, R., Collilieux, X., & Garayt, B. (2012). IGS08:

the IGS realization of ITRF2008. *GPS Solutions*, 16(4), 483–494.

<https://doi.org/10.1007/s10291-011-0248-2>

Revil, A., & Jardani, A. (2013). *The Self-Potential Method: Theory and Applications in Environmental Geosciences*. Cambridge University Press.

Rizzo, E., Suski, B., Revil, A., Straface, S., & Troisi, S. (2004). Self-potential signals

associated with pumping tests experiments. *Journal of Geophysical Research: Solid*

Earth, 109(B10), 14 pp. <https://doi.org/10.1029/2004JB003049>

Rong, Y., Wang, Q., Ding, X., & Huang, Q. (2012). Non-uniform scaling behavior in

Ultra-Low-Frequency (ULF) geomagnetic signals possibly associated with the 2011

M9.0 Tohoku earthquake. *Chinese Journal of Geophysics*, 55(11), 3709–3717.

<https://doi.org/10.6038/j.issn.0001-5733.2012.11.018>

Rundle, J. B., Turcotte, D. L., & Klein, W. (Eds.). (2000). *Geocomplexity and the Physics of Earthquakes*. U. S. A.: American Geophysical Union.

Rydelek, P. A., & Sacks, I. S. (1989). Testing the completeness of earthquake catalogues and

the hypothesis of self-similarity. *Nature*, 337(6204), 251–253.

<https://doi.org/10.1038/337251a0>

Saito, S., Kaida, D., Hattori, K., Febriani, F., & Yoshino, C. (2011). Signal discrimination of ULF electromagnetic data with using singular spectrum analysis – an attempt to detect train noise. *Nat. Hazards Earth Syst. Sci.*, *11*(7), 1863–1874.
<https://doi.org/10.5194/nhess-11-1863-2011>

Sarlis, N. V., Skordas, E. S., Varotsos, P. A., Nagao, T., Kamogawa, M., & Uyeda, S. (2015). Spatiotemporal variations of seismicity before major earthquakes in the Japanese area and their relation with the epicentral locations. *Proceedings of the National Academy of Sciences*, *112*(4), 986–989. <https://doi.org/10.1073/pnas.1422893112>

Scargle, J. D. (1982). Studies in astronomical time series analysis. II. Statistical aspects of spectral analysis of unevenly spaced data. *The Astrophysical Journal*, *263*, 835–853.

Schmidt, J. P., Taylor, R. K., & Gehl, R. J. (2003). Developing Topographic Maps Using a Sub-Meter Accuracy Global Positioning Receiver. *Applied Engineering in Agriculture*, *19*(3), 291–300. <https://doi.org/10.13031/2013.13661>

Schorlemmer, D., Neri, G., Wiemer, S., & Mostaccio, A. (2003). Stability and significance tests for b-value anomalies: Example from the Tyrrhenian Sea. *Geophysical Research Letters*, *30*(16). <https://doi.org/10.1029/2003GL017335>

Seats, K. J., Lawrence, J. F., & Prieto, G. A. (2012). Improved ambient noise correlation functions using Welch's method. *Geophysical Journal International*, *188*(2), 513–523.
<https://doi.org/10.1111/j.1365-246X.2011.05263.x>

Sen, P. K. (1968). Estimates of the Regression Coefficient Based on Kendall's Tau. *Journal of the American Statistical Association*, 63(324), 1379–1389.

<https://doi.org/10.2307/2285891>

Shen, Z.-K., Wang, M., Zeng, Y., & Wang, F. (2015). Optimal Interpolation of Spatially Discretized Geodetic Data. *Bulletin of the Seismological Society of America*, 105(4),

2117–2127. <https://doi.org/10.1785/0120140247>

Shin, T.-C. (1992). Some implications of Taiwan tectonic features from the data collected by the Central Weather Bureau Seismic Network. *Meteorological Bulletin, CWB*, 38(1),

23–48.

Shin, T.-C., Chang, C.-H., Pu, H.-C., Lin, H.-W., & Leu, P.-L. (2013). The Geophysical Database Management System in Taiwan. *Terrestrial, Atmospheric and Oceanic*

Sciences, 24(1), 11–18.

Sibuet, J.-C., Letouzey, J., Barbier, F., Charvet, J., Foucher, J.-P., Hilde, T. W. C., et al.

(1987). Back Arc Extension in the Okinawa Trough. *Journal of Geophysical*

Research: Solid Earth, 92(B13), 14041–14063.

<https://doi.org/10.1029/JB092iB13p14041>

Sill, W. R. (1983). Self-potential modeling from primary flows. *Geophysics*, 48(1), 76–86.

<https://doi.org/10.1190/1.1441409>

Simpson, D. W. (1976). Seismicity changes associated with reservoir loading. *Engineering*

Geology, 10(2), 123–150. [https://doi.org/10.1016/0013-7952\(76\)90016-8](https://doi.org/10.1016/0013-7952(76)90016-8)

Simpson, F., & Bahr, K. (2005). *Practical Magnetotellurics*. Cambridge University Press.

Sornette, A., & Sornette, D. (1990). Earthquake rupture as a critical point: consequences for telluric precursors. *Tectonophysics*, 179(3–4), 327–334. [https://doi.org/10.1016/0040-1951\(90\)90298-M](https://doi.org/10.1016/0040-1951(90)90298-M)

Stumpf, M. P. H., & Porter, M. A. (2012). Critical Truths About Power Laws. *Science*, 335(6069), 665–666. <https://doi.org/10.1126/science.1216142>

Szarka, L. (1988). Geophysical aspects of man-made electromagnetic noise in the earth—A review. *Surveys in Geophysics*, 9(3), 287–318. <https://doi.org/10.1007/BF01901627>

Takahashi, A., Hashimoto, M., Hu, J.-C., Takeuchi, K., Tsai, M.-C., & Fukahata, Y. (2019). Hierarchical Cluster Analysis of Dense GPS Data and Examination of the Nature of the Clusters Associated With Regional Tectonics in Taiwan. *Journal of Geophysical Research: Solid Earth*, 124(5), 5174–5191. <https://doi.org/10.1029/2018JB016995>

Takahashi, I., Hattori, K., Harada, M., Yoshino, C., & Isezaki, N. (2007). Anomalous geoelectrical and geomagnetic signals observed at Southern Boso Peninsula, Japan. *Annals of Geophysics*, 50(1), 123–135.

Telesca, L., Lovallo, M., Romano, G., Konstantinou, K. I., Hsu, H.-L., & Chen, C.-C. (2014). Using the informational Fisher–Shannon method to investigate the influence of long-term deformation processes on geoelectrical signals: An example from the

Taiwan orogeny. *Physica A: Statistical Mechanics and Its Applications*, 414, 340–351.

<https://doi.org/10.1016/j.physa.2014.07.060>

Theil, H. (1992). A Rank-Invariant Method of Linear and Polynomial Regression Analysis.

In B. Raj & J. Koerts (Eds.), *Henri Theil's Contributions to Economics and Econometrics: Econometric Theory and Methodology* (pp. 345–381). Dordrecht:

Springer Netherlands. https://doi.org/10.1007/978-94-011-2546-8_20

Turcotte, D. L. (1999). Seismicity and self-organized criticality. *Physics of the Earth and*

Planetary Interiors, 111(3–4), 275–293.

[https://doi.org/10.1016/S0031-9201\(98\)00167-8](https://doi.org/10.1016/S0031-9201(98)00167-8)

Turcotte, D. L. (1992). Fractals, chaos, self-organized criticality and tectonics. *Terra Nova*,

4(1), 4–12. <https://doi.org/10.1111/j.1365-3121.1992.tb00444.x>

Tylka, J. G., Boren, B. B., & Choueiri, E. Y. (2017). A Generalized Method for

Fractional-Octave Smoothing of Transfer Functions that Preserves Log-Frequency Symmetry. *Journal of the Audio Engineering Society*, 65(3), 239–245.

Uyeda, S., Hayakawa, M., Nagao, T., Molchanov, O., Hattori, K., Orihara, Y., et al. (2002).

Electric and magnetic phenomena observed before the volcano-seismic activity in 2000 in the Izu Island Region, Japan. *Proceedings of the National Academy of Sciences*, 99(11), 7352–7355. <https://doi.org/10.1073/pnas.072208499>

Varotsos, P. A., Sarlis, N. V., & Skordas, E. S. (2003). Long-range correlations in the electric

signals that precede rupture: Further investigations. *Physical Review E*, 67(2), 021109.

<https://doi.org/10.1103/PhysRevE.67.021109>

Viljanen, A., & Pirjola, R. (1994). Geomagnetically induced currents in the Finnish high-voltage power system. *Surveys in Geophysics*, 15(4), 383–408.

<https://doi.org/10.1007/BF00665999>

Virkar, Y., & Clauset, A. (2014). Power-law distributions in binned empirical data. *The Annals of Applied Statistics*, 8(1), 89–119. <https://doi.org/10.1214/13-AOAS710>

Wang, J. H. (1989). The Taiwan telemetered seismographic network. *Physics of the Earth and Planetary Interiors*, 58(1), 9–18. [https://doi.org/10.1016/0031-9201\(89\)90090-3](https://doi.org/10.1016/0031-9201(89)90090-3)

Welch, P. (1967). The use of fast Fourier transform for the estimation of power spectra: A method based on time averaging over short, modified periodograms. *IEEE Transactions on Audio and Electroacoustics*, 15(2), 70–73.

<https://doi.org/10.1109/TAU.1967.1161901>

Wiemer, S., & McNutt, S. R. (1997). Variations in the frequency-magnitude distribution with depth in two volcanic areas: Mount St. Helens, Washington, and Mt. Spurr, Alaska. *Geophysical Research Letters*, 24(2), 189–192. <https://doi.org/10.1029/96GL03779>

Wing, M. G., & Frank, J. (2011). Vertical measurement accuracy and reliability of mapping-grade GPS receivers. *Computers and Electronics in Agriculture*, 78(2), 188–194. <https://doi.org/10.1016/j.compag.2011.07.006>

Accepted Article

Wu, Y.-M., Chang, C.-H., Zhao, L., Teng, T.-L., & Nakamura, M. (2008). A Comprehensive Relocation of Earthquakes in Taiwan from 1991 to 2005. *Bulletin of the Seismological Society of America*, 98(3), 1471–1481.

<https://doi.org/10.1785/0120070166>

Wyss, M., Shimazaki, K., & Wiemer, S. (1997). Mapping active magma chambers by b values beneath the off-Ito volcano, Japan. *Journal of Geophysical Research: Solid Earth*, 102(B9), 20413–20422. <https://doi.org/10.1029/97JB01074>

Xu, G., Han, P., Huang, Q., Hattori, K., Febriani, F., & Yamaguchi, H. (2013). Anomalous behaviors of geomagnetic diurnal variations prior to the 2011 off the Pacific coast of Tohoku earthquake (Mw9.0). *Journal of Asian Earth Sciences*, 77, 59–65.

<https://doi.org/10.1016/j.jseaes.2013.08.011>

Yu, S.-B., Kuo, L., Punongbayan, R. S., & Ramos, E. G. (1999). GPS observation of crustal deformation in the Taiwan-Luzon Region. *Geophysical Research Letters*, 26(7), 923–926. <https://doi.org/10.1029/1999GL900148>

Yu, S.-B., & Kuo, L.-C. (2001). Present-day crustal motion along the Longitudinal Valley Fault, eastern Taiwan. *Tectonophysics*, 333(1), 199–217.

[https://doi.org/10.1016/S0040-1951\(00\)00275-4](https://doi.org/10.1016/S0040-1951(00)00275-4)

Yu, S.-B., Chen, H.-Y., & Kuo, L.-C. (1997). Velocity field of GPS stations in the Taiwan area. *Tectonophysics*, 274(1), 41–59. [https://doi.org/10.1016/S0040-1951\(96\)00297-1](https://doi.org/10.1016/S0040-1951(96)00297-1)

Zhang, X., Ye, Q., & Liu, G. (2019). Discussion on the sea–land telluric current vector and its continuity during a geomagnetic storm based on coastal stations in China and Canada. *Geomatics, Natural Hazards and Risk*, 10(1), 1805–1821.

<https://doi.org/10.1080/19475705.2019.1629184>

Zumberge, J. F., Heflin, M. B., Jefferson, D. C., Watkins, M. M., & Webb, F. H. (1997). Precise point positioning for the efficient and robust analysis of GPS data from large networks. *Journal of Geophysical Research: Solid Earth*, 102(B3), 5005–5017.

<https://doi.org/10.1029/96JB03860>

Table 1. Properties of the SP stations on the GEMS network. The first three columns are the name, location, and the start date of a station, respectively. The fourth column is the number of the segments for estimating the PSDs using the WD, DT, and NT data. The fifth column is the urbanization level (U) estimated by Huang et al. (2018). The sixth column is the slope of the Gutenberg-Richter distribution (b) estimated by the catalog selected during 2012–2017 and within a 50-km radius of each SP station. The seventh column is the dilation strain rate (D) in $\mu\text{strain/yr}$ derived from the GPS data during 2012–2017. The b and D values are represented by mean within one standard deviation.

Station Name	Location ($^{\circ}\text{E}$, $^{\circ}\text{N}$)	Start Date (yyyy-mm-dd)	Number of Analyzed Segments (WD, DT, NT)	U	b	D ($\mu\text{strain/yr}$)
SHRL	121.56, 25.16	2012-05-02	2012, 2025, 2025	4	1.74 \pm 0.29	0.0783 \pm 0.0580
KUOL	121.14, 24.96	2012-01-01	2133, 2144, 2147	3*	1.22 \pm 0.05	-0.0614 \pm 0.0507
TOCH	121.81, 24.84	2013-08-29	1356, 1365, 1367	3	1.14 \pm 0.03	0.5416 \pm 0.0395
HUAL	121.37, 24.67	2012-01-17	1710, 1983, 1690	1	1.07 \pm 0.02	0.0025 \pm 0.0301
ENAN	121.79, 24.48	2012-02-15	1996, 2024, 2020	3	0.94 \pm 0.02	0.2333 \pm 0.0520
DAHU	120.90, 24.41	2012-02-07	2117, 2125, 2125	1	1.05 \pm 0.03	-0.1885 \pm 0.0348
LISH	121.25, 24.25	2012-11-07	1717, 1744, 1731	1	0.92 \pm 0.03	0.0586 \pm 0.0321
SHCH	121.63, 24.12	2012-04-23	1997, 2009, 2004	3	0.90 \pm 0.01	-0.3252 \pm 0.0593
HERM	120.50, 24.11	2012-02-09	2055, 2071, 2066	2	0.99 \pm 0.04	-0.0984 \pm 0.0481
PULI	120.98, 23.92	2012-03-01	2022, 2027, 2026	2*	0.88 \pm 0.04	-0.4940 \pm 0.0416
FENL	121.41, 23.72	2012-04-23	1953, 1968, 1970	1	0.80 \pm 0.01	-0.7178 \pm 0.0429
SIHU	120.23, 23.64	2012-02-08	1958, 2001, 1991	1	1.05 \pm 0.04	-0.1375 \pm 0.0490
DABA	120.75, 23.45	2013-02-01	1772, 1772, 1773	3	0.90 \pm 0.01	-0.4715 \pm 0.0446
YULI	121.32, 23.32	2012-04-23	2024, 2037, 2035	3	0.85 \pm 0.01	-1.6821 \pm 0.0506
CHCH	120.16, 23.22	2012-03-01	2019, 2031, 2025	2	0.99 \pm 0.03	-0.2191 \pm 0.0396
LIOQ	120.66, 23.03	2012-07-20	1943, 1945, 1941	1	0.88 \pm 0.01	-0.2529 \pm 0.0543
RUEY	121.16, 22.97	2012-04-23	2044, 2052, 2052	1	0.86 \pm 0.01	-1.7620 \pm 0.0621
KAOH	120.29, 22.66	2012-02-23	2067, 2073, 2069	4	0.95 \pm 0.03	-0.0944 \pm 0.0658
WANL	120.59, 22.59	2012-12-20	1681, 1703, 1702	3	0.94 \pm 0.02	-0.2709 \pm 0.0410
FENG	120.70, 22.20	2012-12-20	1400, 1528, 1493	1	0.98 \pm 0.03	0.4615 \pm 0.0347

*Note that (1) the urbanization level at KUOL is set $U = 3$ due to its proximity between areas

of $U = 2$ and $U = 4$; (2) similarly, we set $U = 2$ at PULI between areas of $U = 1$ and $U = 3$.

Table 2. Mean values and standard deviations of the SP power-law parameters estimated by fitting Eq. (2) to the PSDs between 0.001 and 0.1 Hz by using the WD, DT, and NT data for the NS and EW components.

Station Name	β_{WD} (NS; EW)	β_{DT} (NS; EW)	β_{NT} (NS; EW)	α_{WD} (NS; EW)	α_{DT} (NS; EW)	α_{NT} (NS; EW)
SHRL	1.14±0.51; 1.20±0.38	1.33±0.48; 1.20±0.44	1.19±0.42; 1.33±0.38	-0.92±1.09; -0.17±0.99	-1.42±1.12; -0.64±1.14	-1.42±0.92; -0.73±0.97
KUOL	1.22±0.31; 1.26±0.28	1.16±0.30; 1.23±0.27	1.22±0.33; 1.27±0.32	-1.32±0.91; -1.80±1.31	-1.27±0.93; -1.82±1.28	-1.72±1.09; -2.24±1.33
TOCH	1.15±0.30; 1.17±0.21	1.20±0.32; 1.17±0.23	1.28±0.33; 1.24±0.27	-0.97±1.02; -0.34±0.93	-1.05±1.10; -0.30±1.04	-2.01±1.10; -1.41±0.80
HUAL	1.58±0.21; 1.41±0.32	1.52±0.22; 1.44±0.25	1.49±0.29; 1.38±0.34	0.21±0.93; -0.15±1.14	0.35±1.23; -0.48±0.90	0.07±0.99; -0.27±1.32
ENAN	1.02±0.28; 0.87±0.26	1.20±0.27; 1.15±0.28	1.08±0.35; 1.01±0.33	-0.18±0.61; 0.67±0.66	-0.81±0.55; -0.46±0.65	-0.61±0.96; -0.35±1.31
DAHU	1.46±0.20; 1.45±0.23	1.40±0.24; 1.38±0.26	1.54±0.23; 1.58±0.29	-0.50±0.85; -0.33±1.01	-0.46±0.98; -0.24±1.06	-1.00±1.02; -0.95±1.18
LISH	1.36±0.21; 1.31±0.25	1.44±0.26; 1.43±0.26	1.29±0.27; 1.24±0.30	-0.91±0.68; -0.26±0.72	-1.25±1.21; -0.75±1.28	-1.06±0.88; -0.44±1.08
SHCH	1.10±0.18; 1.07±0.23	1.13±0.23; 1.04±0.27	0.96±0.22; 0.97±0.25	-0.89±0.86; -1.00±0.77	-1.11±0.96; -1.15±0.93	-0.87±0.86; -1.02±0.72
HERM	1.06±0.23; 1.05±0.24	0.98±0.27; 1.01±0.27	0.93±0.22; 0.95±0.24	-0.60±0.35; -0.14±0.37	-0.72±0.37; -0.40±0.39	-0.55±0.46; -0.10±0.53
PULI	1.22±0.22; 1.15±0.27	1.16±0.27; 1.19±0.32	1.17±0.23; 1.14±0.23	-0.51±0.62; -0.89±0.65	-0.98±0.53; -1.51±0.56	-0.63±0.72; -1.12±0.80
FENL	1.36±0.26; 1.15±0.33	1.30±0.28; 1.06±0.33	1.26±0.33; 1.11±0.33	-0.06±0.42; 0.19±0.78	-0.02±0.49; 0.04±0.70	-0.55±0.54; -0.12±0.98
SIHU	1.15±0.28; 1.14±0.27	1.24±0.30; 1.24±0.30	1.12±0.29; 1.12±0.29	-0.82±1.88; -0.21±1.25	-1.52±2.09; -0.92±1.52	-0.71±1.96; -0.20±1.38
DABA	1.22±0.23; 1.24±0.27	1.19±0.31; 1.21±0.31	1.13±0.26; 1.23±0.33	-0.34±0.46; -0.39±0.38	-0.46±0.51; -0.38±0.43	-0.44±0.58; -0.72±0.48
YULI	1.18±0.27; 0.93±0.37	1.19±0.29; 1.06±0.38	1.11±0.25; 1.03±0.31	-1.50±0.55; -0.71±0.67	-1.52±0.60; -1.05±0.75	-2.03±0.50; -1.40±0.57
CHCH	1.48±0.26; 1.42±0.31	1.50±0.27; 1.46±0.27	1.32±0.25; 1.38±0.31	-2.50±0.43; -2.51±0.59	-2.86±0.53; -2.83±0.63	-2.52±0.54; -2.79±0.52
LIOQ	1.35±0.24; 1.36±0.23	1.39±0.26; 1.40±0.27	1.33±0.23; 1.32±0.24	-0.22±0.69; -0.54±0.75	-0.40±0.75; -0.79±0.81	-0.33±0.74; -0.67±0.74
RUEY	1.17±0.23; 1.49±0.23	1.10±0.25; 1.59±0.28	1.15±0.24; 1.44±0.29	-0.61±0.36; -1.32±0.54	-0.80±0.49; -1.70±0.73	-0.77±0.40; -1.53±0.52
KAOH	0.80±0.26; 0.87±0.25	0.96±0.37; 0.92±0.33	0.96±0.21; 0.88±0.20	-0.97±0.51; -1.20±0.53	-1.09±1.53; -1.27±1.23	-2.25±0.48; -2.42±0.41
WANL	1.28±0.36; 1.13±0.29	1.41±0.36; 1.43±0.33	1.12±0.30; 1.09±0.26	-1.34±0.55; -1.49±0.62	-1.92±0.67; -2.16±0.71	-1.67±0.72; -1.76±0.88
FENG	1.76±0.20; 1.37±0.24	1.86±0.21; 1.48±0.26	1.64±0.28; 1.36±0.30	-2.53±0.37; -2.15±0.52	-2.78±0.44; -2.54±0.56	-2.53±0.49; -2.31±0.61

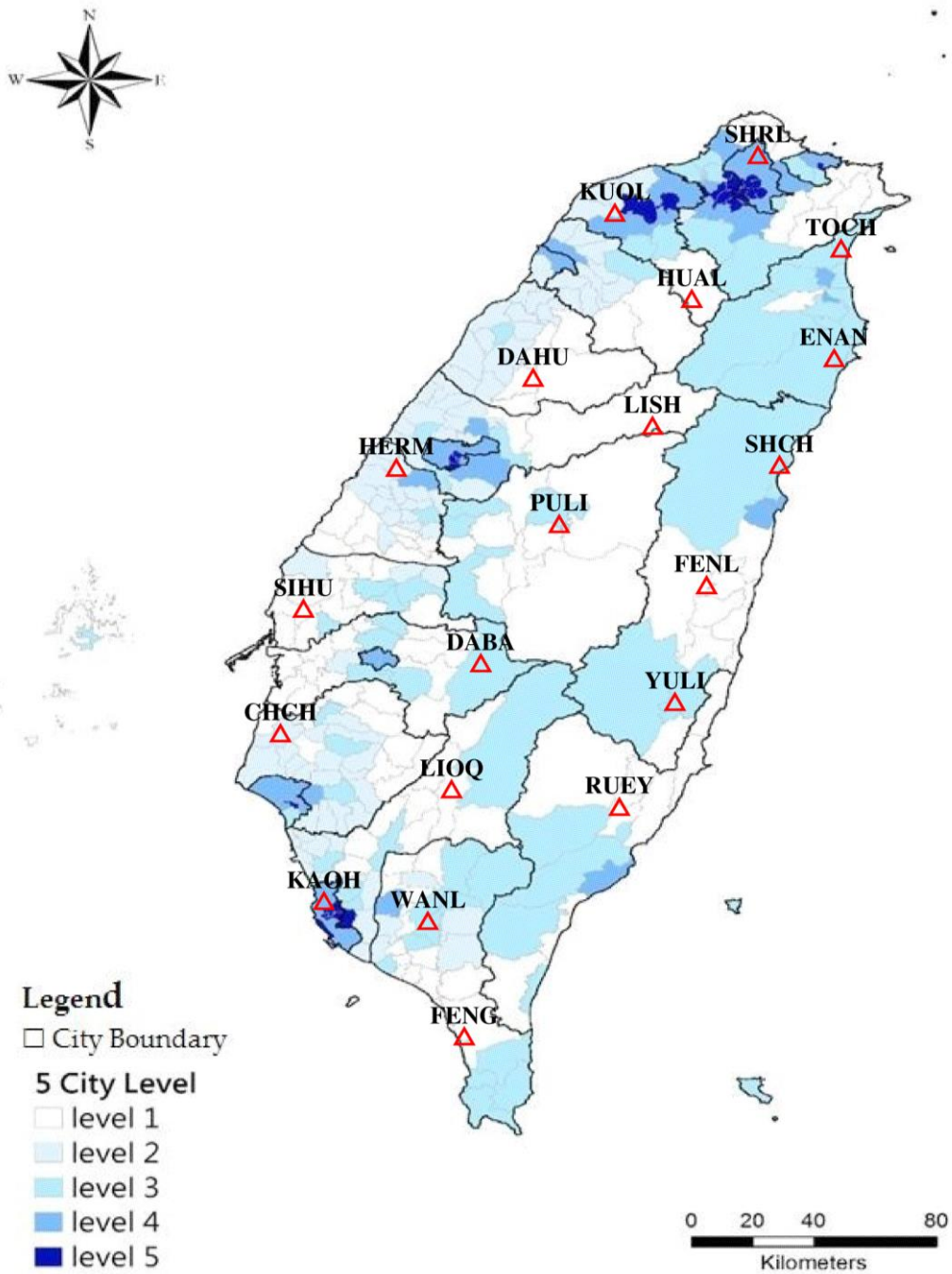


Figure 1. Maps of the SP stations on the GEMS network (red triangles) and of the urbanization levels of the Taiwan townships (blue shading) as estimated by Huang et al. (2018).

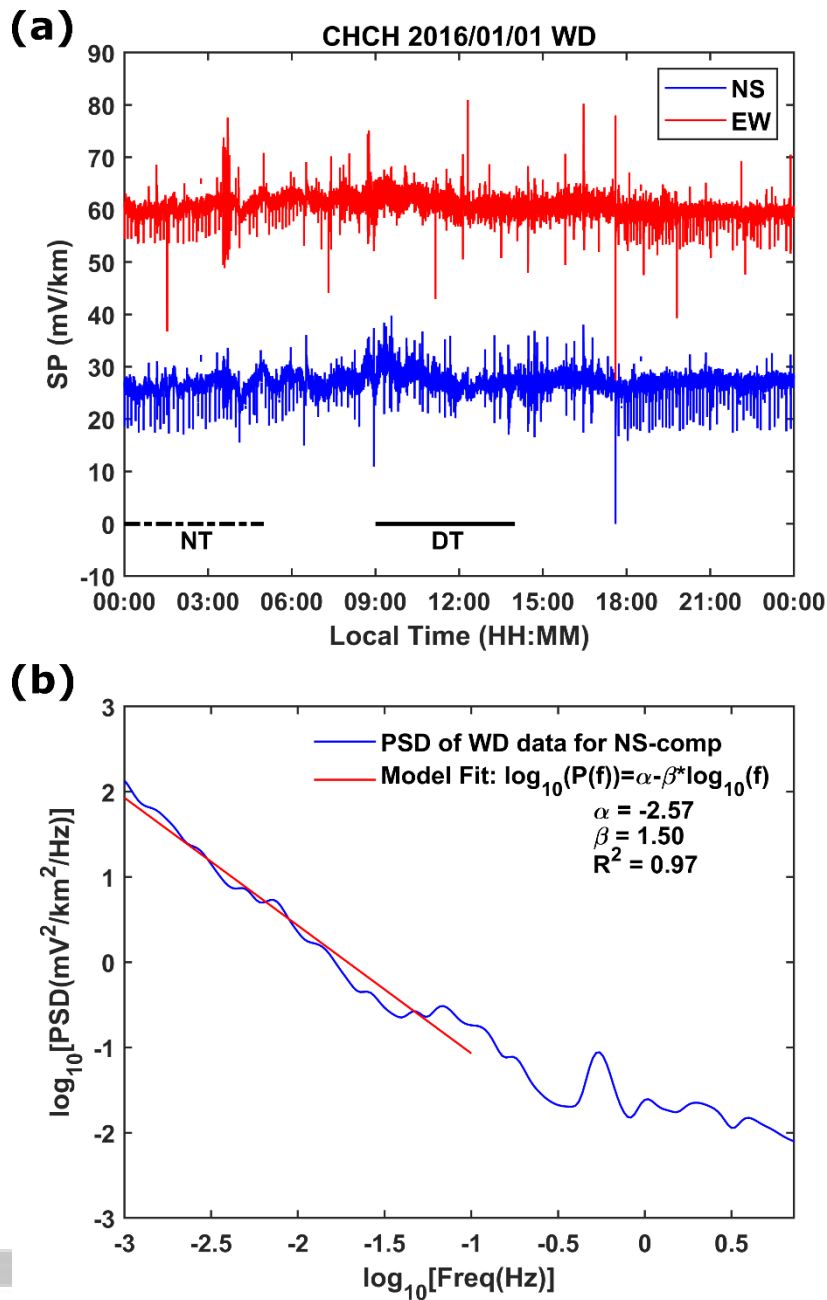


Figure 2. (a) SP time series on 1 January 2016 at CHCH for the NS (blue) and EW (red) components. The full-time scale represents the whole day (00:00–24:00 Local Time, WD), the black dashed line represents the nighttime (00:00–05:00 Local Time, NT), and the black solid line represents the daytime (09:00–14:00 Local Time, DT). (b) Illustration of the power spectral density estimated using the WD data for the NS component in (a) and the fitting of Eq. (2) (see Section 3.4) to the PSDs between 0.001 and 0.1 Hz by using the least-squares regression in a log-log scale.

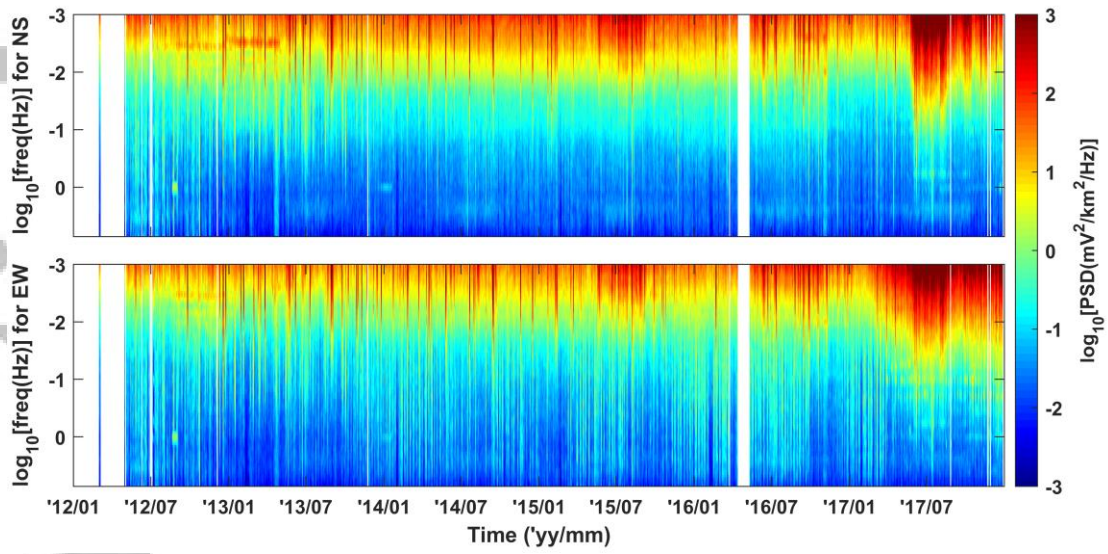


Figure 3. Spectrogram of the SP signals using the WD data at CHCH for the NS (top) and EW (bottom) components. The white areas correspond to discarded SP data or data missings.

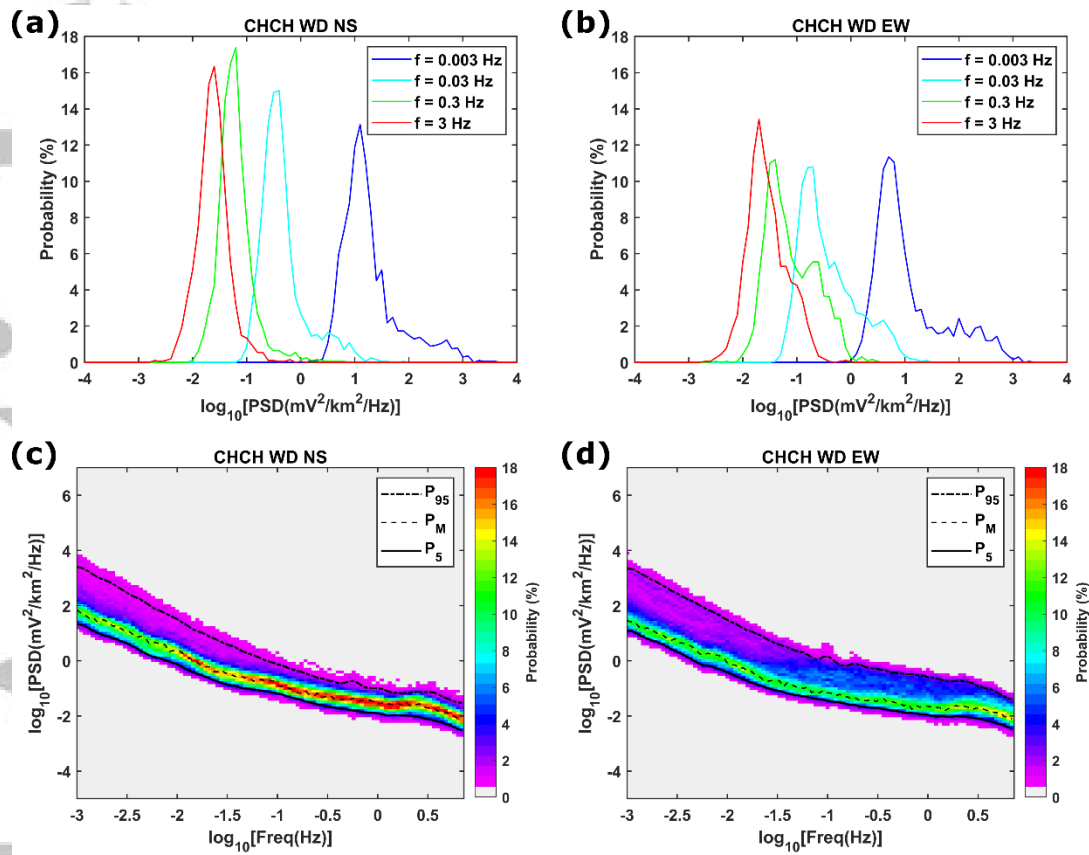


Figure 4. Probability density functions (PDFs) at CHCH using the WD data at frequencies of 0.003, 0.03, 0.3, and 3 Hz for the (a) NS and (b) EW components. Probability power spectral densities (PPSDs) at CHCH using the WD data for the (c) NS and (d) EW components. The PPSD shows the PDF at each frequency in one plot. The P_{95} , P_M , and P_5 models represent the 95th percentile, mode, and 5th percentile of the PDFs.

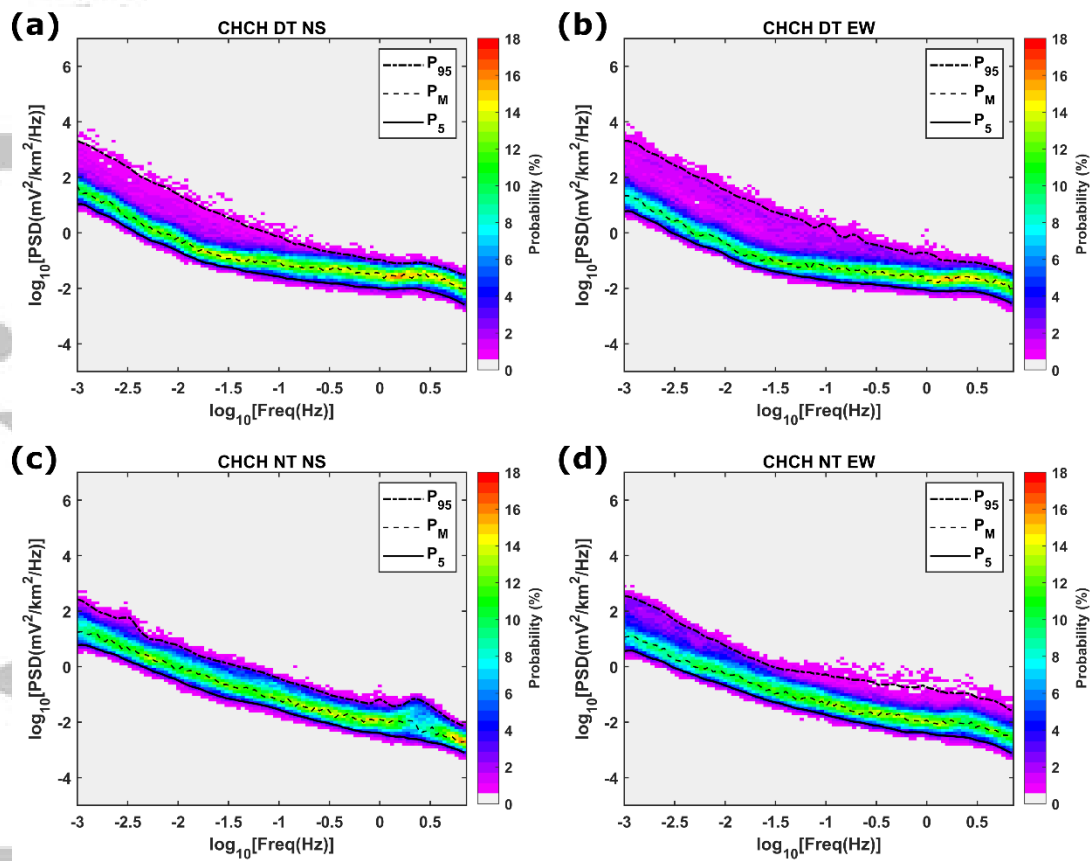


Figure 5. Probability power spectral densities (PPSDs) at CHCH using the DT data for the (a) NS and (b) EW components. PPSDs at CHCH using the NT data for the (c) NS and (d) EW components.

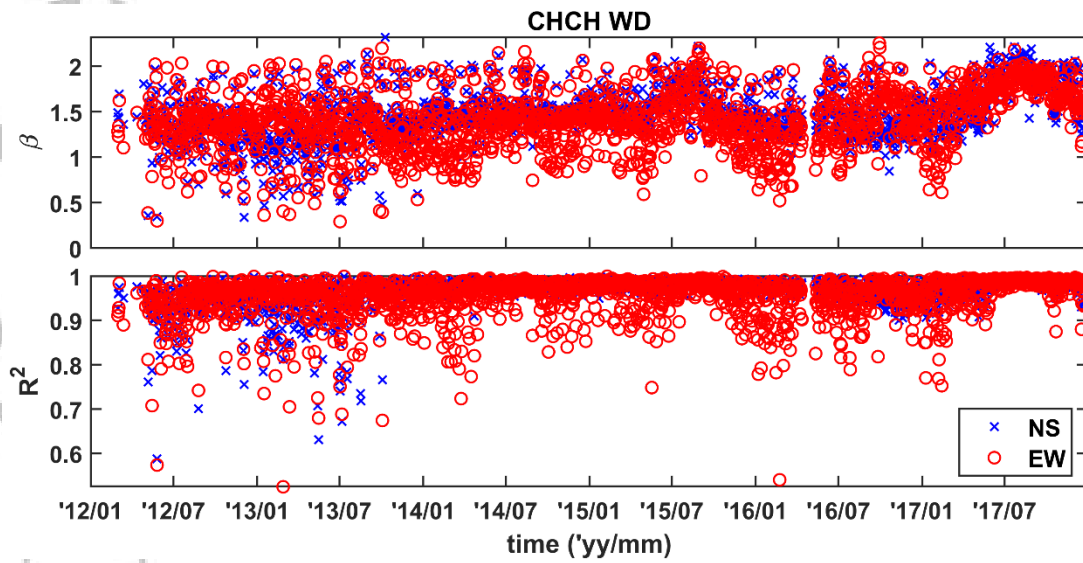


Figure 6. Time series of the spectral scaling of the SP $1/f^\beta$ noises (top) and coefficient of determination R^2 (bottom) at CHCH for the WD data by fitting the PSDs between 0.001 and 0.1 Hz.

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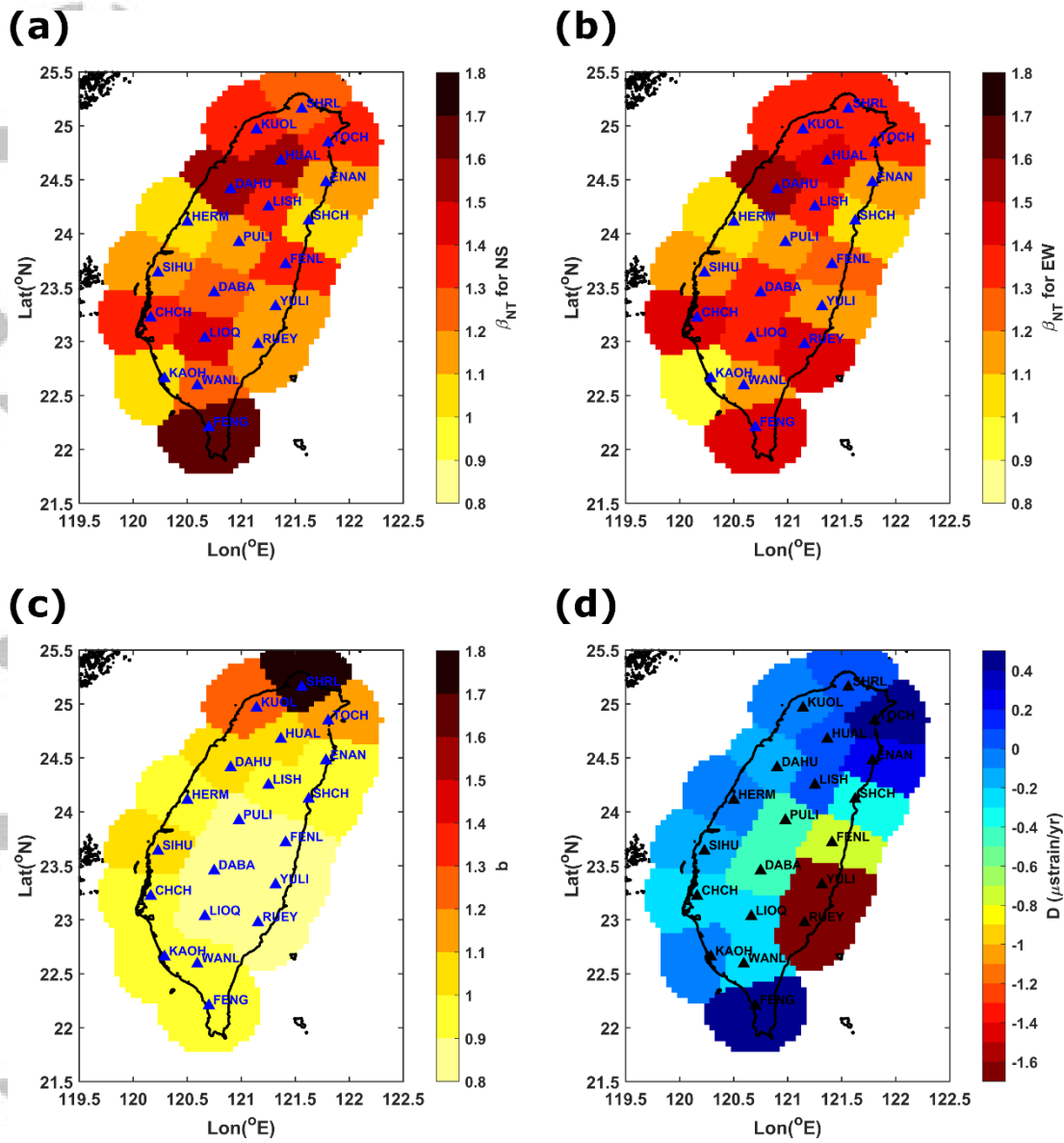


Figure 7. Maps of the SP scalings $\bar{\beta}_{NT}$ listed in Table 2 for the (a) NS and (b) EW components. Maps of (c) the GR \bar{b} values and (d) the dilation strain rates (\bar{D}) listed in Table 1.

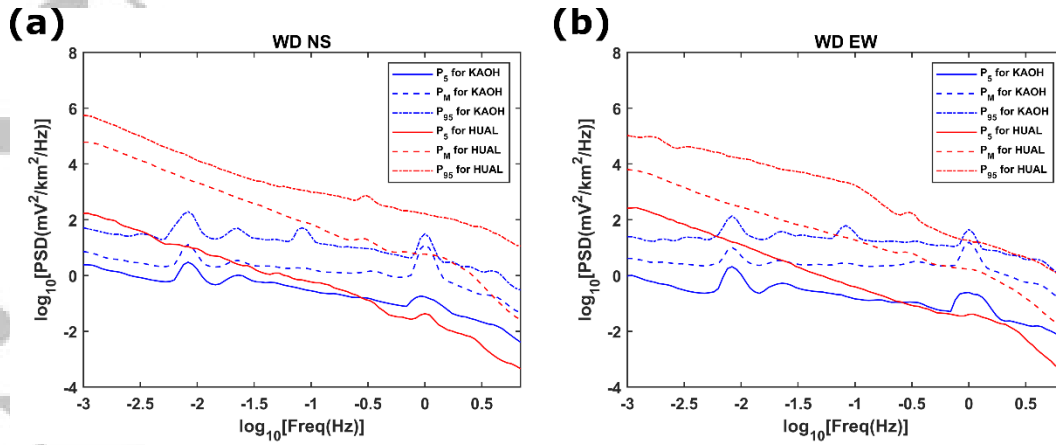


Figure 8. Ambient SP noise models (P_5 , P_M , and P_{95}) at KAOH (blue) and HUAL (red) using the WD data for the (a) NS and (b) EW components.

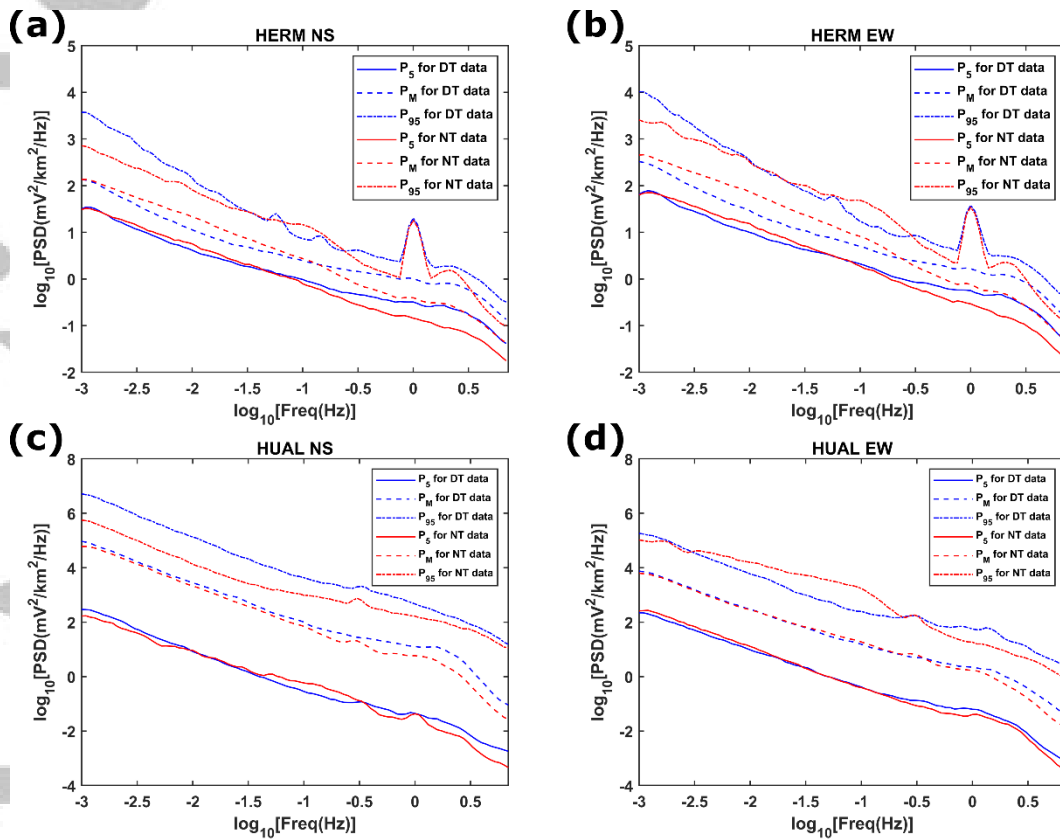


Figure 9. Ambient SP noise models (P_5 , P_M , and P_{95}) at HERM using the DT and NT data for the (a) NS and (b) EW components. Similarly, noise models at HUAL for the (c) NS and (d) EW components.

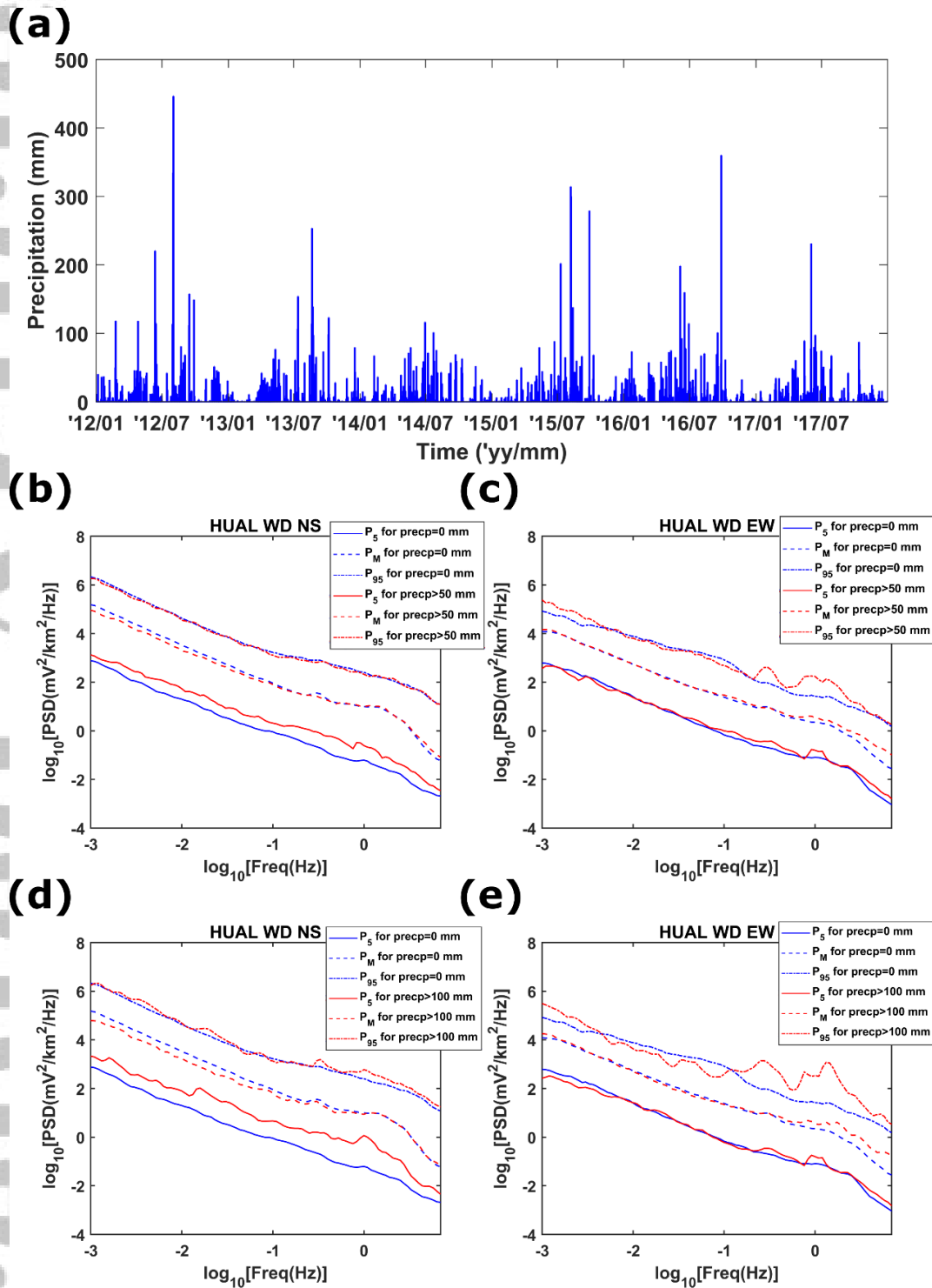


Figure 10. (a) Daily time series of the precipitation at HUAL during 2012–2017. Ambient SP noise models (P_5 , P_M , and P_{95}) when precipitation is 0 mm (blue) and greater than 50 mm (red) at HUAL using the WD data for the (a) NS and (b) EW components. Similarly, noise models when precipitation is 0 mm (blue) and greater than 100 mm (red) for the (a) NS and (b) EW components.

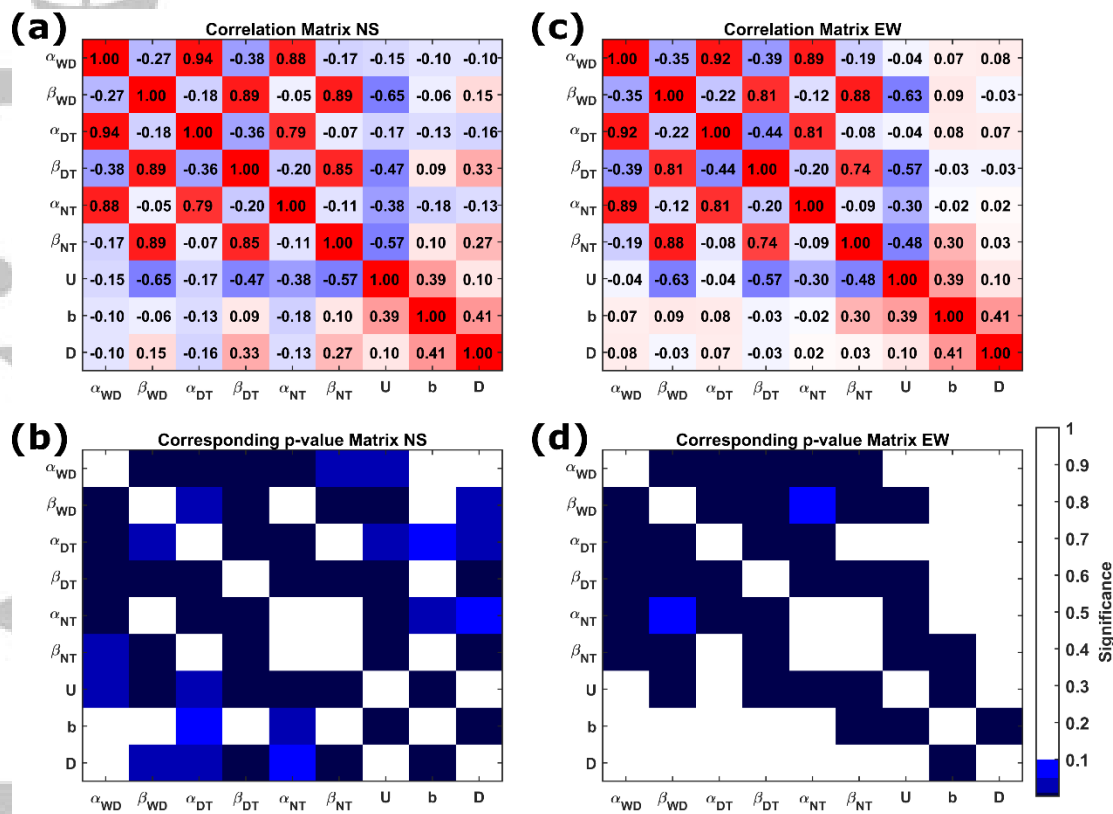


Figure 11. Correlation matrix among the urbanization level (U), GR \bar{b} value, dilation strain rate (\bar{D}) and the SP power-law parameters ($\bar{\alpha}$ and $\bar{\beta}$) estimated between 0.001 and 0.1 Hz, as an example, for the (a) NS and (b) EW components. Corresponding p-value matrices for the (a) NS and (b) EW components. There are three significance levels of 0.1, 0.05, and 0.001 for the color bar.

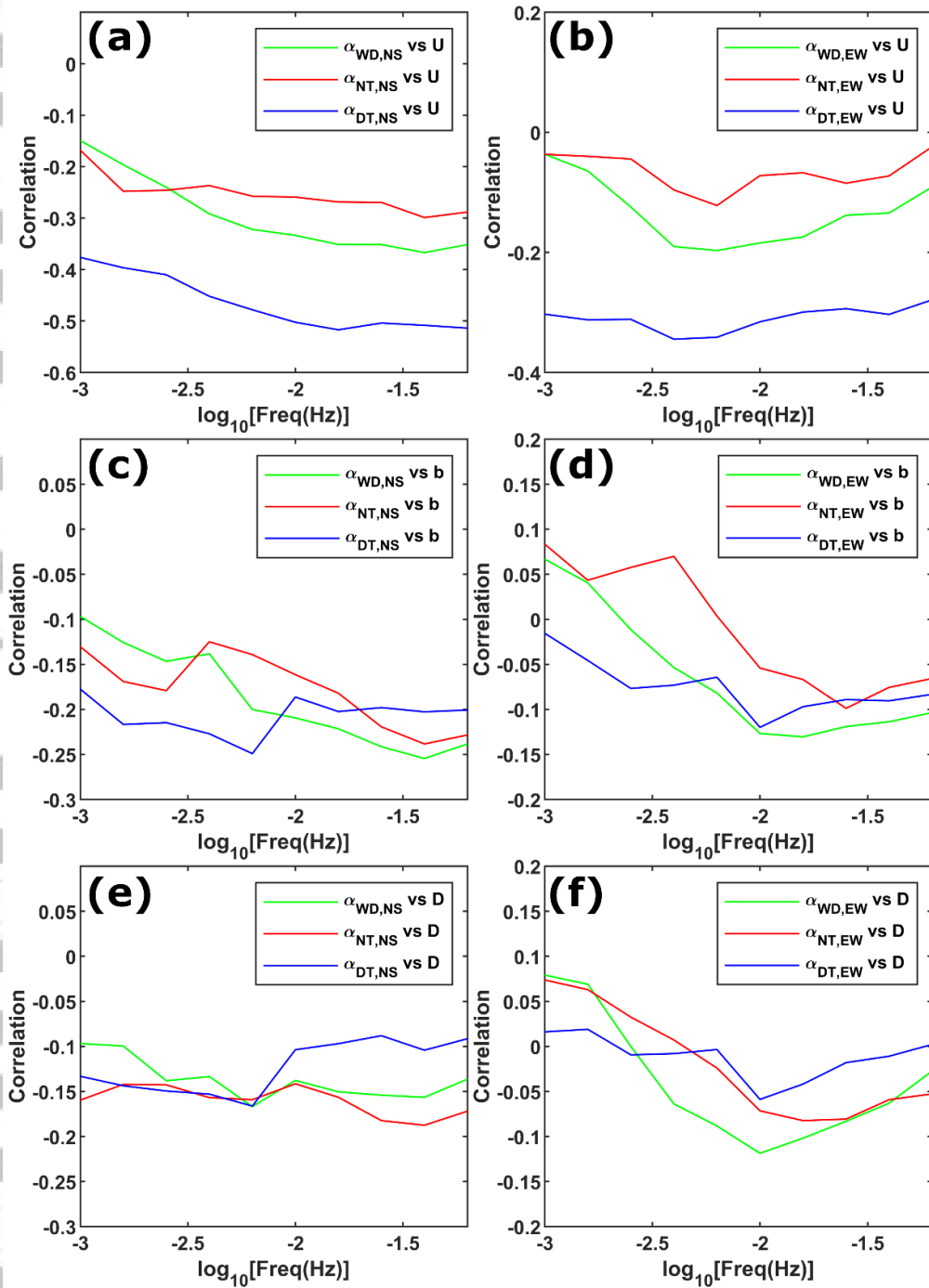


Figure 12. Correlations between U and $\bar{\alpha}$ estimated using the different time-spanned data at different frequency bands for the (a) NS and (b) EW components. Similarly, correlations between \bar{b} and $\bar{\alpha}$ for the (c) NS and (d) EW components and correlations between \bar{D} and $\bar{\alpha}$ for the (e) NS and (f) EW components. The frequency tick in a logarithmic scale here represents the frequency band between 10^{tick} and $10^{\text{tick}+2}$ Hz (see Section 3.4).

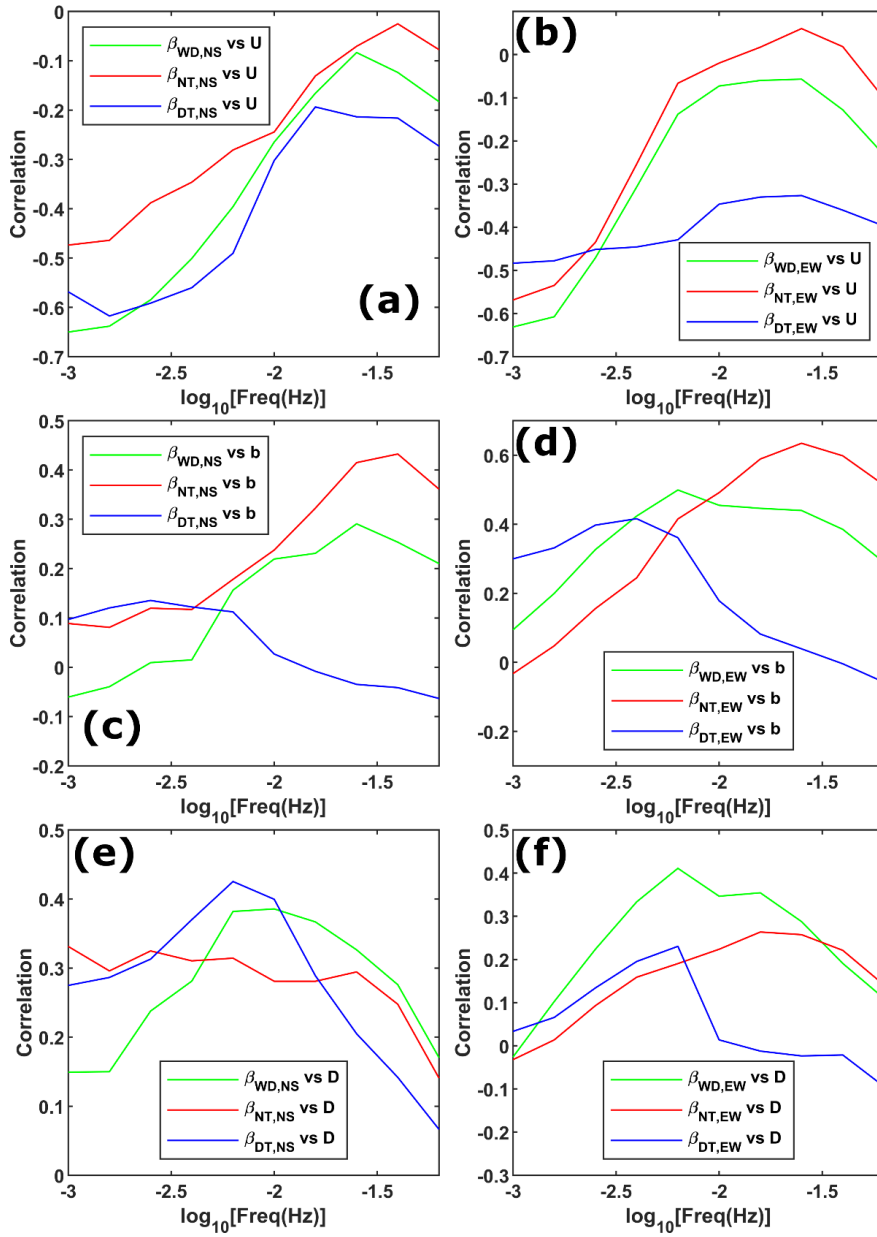


Figure 13. Correlations between U and $\bar{\beta}$ estimated using the different time-spanned data at different frequency bands for the (a) NS and (b) EW components. Similarly, correlations between \bar{b} and $\bar{\beta}$ for the (c) NS and (d) EW components and correlations between \bar{D} and $\bar{\beta}$ for the (e) NS and (f) EW components. The frequency tick in a logarithmic scale here represents the frequency band between 10^{tick} and $10^{\text{tick}+2}$ Hz (see Section 3.4).