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Received 18 December 2022
ACCEPTED 07 September 2023
published 22 September 2023

## CITATION

Shen J, Wang J, Yang F and An R (2023) Impact of soda tax on beverage price, sale, purchase, and consumption in the US: a systematic review and meta-analysis of natural experiments.
Front. Public Health 11:1126569.
doi: 10.3389/fpubh.2023.1126569

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# Impact of soda tax on beverage price, sale, purchase, and consumption in the US: a systematic review and meta-analysis of natural experiments 

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#### Abstract

Background: As a primary source of added sugars in the US diet, sugar-sweetened beverage (SSB) consumption is presumed to contribute to obesity prevalence and poor oral health. We systematically synthesized and quantified evidence from US-based natural experiments concerning the impact of SSB taxes on beverage prices, sales, purchases, and consumption. Methods: A keyword and reference search was performed in PubMed, Web of Science, Cochrane Library, Scopus, and EconLit from the inception of an electronic bibliographic database to Oct 31, 2022. Meta-analysis was conducted to estimate the pooled effect of soda taxes on SSB consumption, prices, passthrough rate, and purchases. Results: Twenty-six natural experiments, all adopting a difference-in-differences approach, were included. Studies assessed soda taxes in Berkeley, Oakland, and San Francisco in California, Philadelphia in Pennsylvania, Boulder in Colorado, Seattle in Washington, and Cook County in Illinois. Tax rates ranged from 1 to 2 ф/oz. The imposition of the soda tax was associated with a 1.06 ф/oz. $195 \%$ confidence interval $[C I]=0.90,1.22$ ) increase in SSB prices and a $27.3 \%$ ( $95 \%$ $\mathrm{Cl}=19.3,35.4 \%$ ) decrease in SSB purchases. The soda tax passthrough rate was $79.7 \%$ ( $95 \% \mathrm{Cl}=65.8,93.6 \%$ ). A $1 \not \subset / \mathrm{oz}$. increase in soda tax rate was associated with increased prices of SSBs by 0.84 ¢/oz ( $95 \% \mathrm{Cl}=0.33,1.35$ ). Conclusion: Soda taxes could be effective policy leverage to nudge people toward purchasing and consuming fewer SSBs. Future research should examine evidence-based classifications of SSBs, targeted use of revenues generated by taxes to reduce health and income disparities, and the feasibility of redesigning the soda tax to improve efficiency.


## KEYWORDS

sugar-sweetened beverages, tax, systematic review, meta-analysis, natural experiments

## 1. Introduction

Sugar-sweetened beverages (SSBs) are beverages sweetened with various forms of added sugars, including brown sugar, corn syrup, glucose, lactose, and sucrose (1). In the United States, SSB consumption is prevalent, with six in ten youths and five in ten adults consuming SSBs on any given day during the period of 2011$2014(2,3)$. This consumption amounts to over 140 kcal from SSBs per day for both youth and adults $(2,3)$. Extensive epidemiologic studies have consistently documented that SSB consumption, as a primary source of added sugars in the US diet, is a significant contributor to the prevalence of obesity, cardiometabolic diseases, and oral health risk (4-6). These associations have also been observed globally $(7,8)$.

Recent policy interventions aimed at combating obesity have utilized economic incentives to "nudge" (i.e., promote or encourage) people toward a healthier diet choice ( $9-11$ ). For example, healthy food subsidies promote fruit and vegetable intake, while excise taxes aim to discourage the consumption of less desirable foods and beverages. Excise taxes, such as soda taxes, are implemented in countries like the US, France, New Zealand, Netherlands, and South Africa (12). A soda tax is a specific excise tax charged on the sale of SSBs to reduce consumption (13). Merchants pay the tax, which is then passed on to consumers through higher prices. The amount of the tax varies across different regions, as it is applied by both state and federal governments.

Preliminary evidence suggests that soda taxes are associated with weight loss, reduced body mass index (BMI), and decreased risks of overweight and obesity $(14,15)$. Additionally, soda taxes can help address dental health issues and reduce the prevalence of tooth decay (16). Berkeley, California, became the first city in the US to implement a soda tax, which imposed a $1 \phi /$ oz. tax on the distributors of specific SSBs, including soda and sports/energy drinks (17). This tax took effective on January 1, 2015 (17). Following Berkeley's example, several other cities, such as Philadelphia, Pennsylvania (1.5 ф/oz.; 1/1/2017), Albany, California (1 ¢/oz.; 4/1/2017), Oakland, California ( 1 ¢/oz.; 7/1/2017), Boulder, Colorado ( $2 \phi /$ oz.; 7/1/2017), Cook County, Illinois ( 1 ¢/oz.; 8/2/2017), San Francisco, California ( 1 ¢/oz.; 1/1/2018), and Seattle, Washington ( 1.75 ¢/oz.; 1/1/2018), have also implemented soda taxes (18).

Existing research on the effect of soda taxes in the US can be categorized into three groups: "proxy," "modeling," and "local" studies. The "proxy" studies have used state sales tax on soda, candy, and other qualified groceries as a substitute for a soda tax due to the absence of soda taxes or related data (19-24). However, using such a "proxy" can be problematic because a soda tax, as an excise tax, is fundamentally different from a state general or selective sales tax applied as a percentage of the purchase price (24). The "modeling" studies have employed systems science models to simulate the effect of a soda tax using pre-specified parameter values and statistical distributions (25-34). However, these studies differ from other categories of research in that they are prospective in nature. They aim to identify potential effects of policies that have yet to be enacted and, by necessity, make assumptions about possible retailer and consumer responses. Although the "modeling" studies compare simulated counterfactuals to status quo baselines, they do not provide direct causal inference due to their non-experimental study design. The
"local" studies, on the other hand, have utilized quasi-experimental methods to compared soda prices, sales, purchases, or consumption between cities that have implemented a soda tax and neighboring cities without such a tax, or before and after the implementation of a soda tax in a city. Unlike "proxy" or "modeling" studies, the "local" studies directly estimated the impact of soda tax using temporal and geographical variations in tax implementation (3552). Thus, these "local" studies serve as natural experiments that provide valuable causal inference. Powell et al. conducted a review of seven local SSB tax implementations and reported an average tax pass-through rate of $70 \%$ (53). Another review by Powell et al. found that the demand for SSBs declined by $20 \%$, with an estimated price elasticity of demand of -1.5 following the implementation of soda taxes (54).

This study aims to systematically synthesize and quantify the impacts of soda taxes on beverage prices, sales, purchases, and consumption in the US. By exclusively relying on natural experiments (i.e., "local" studies), our review provides robust causal inferences regarding the effectiveness of soda taxes. Moreover, our approach goes beyond a narrative review by providing quantitative estimates of the magnitude of the tax effect. The findings of this study can inform local, state, and federal policymakers in designing or revising soda tax-related legislation and implementation strategies to effectively combat obesity.

## 2. Methods

The present study was conducted following the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (55).

### 2.1. Eligibility criteria

Studies meeting all of the following criteria were eligible for the review: (1) Participants: consumers, stores, or beverage items within US taxing jurisdictions that implemented a soda tax; (2) Interventions: soda tax (SSB excise tax); (3) Comparisons: consumers, stores, or beverage items within and outside the US taxing jurisdictions that implemented a soda tax; (4) Outcomes: beverage prices, sales, purchases, and consumption; (5) Study design: natural experiment; (6) Article type: peer-reviewed original study; (7) Time window of search: from the inception of an electronic bibliographic database to Oct 31, 2022; and (8) Language: English.

### 2.2. Search strategy

A keyword search was performed in five electronic bibliographic databases: PubMed, Web of Science, Cochrane Library, Scopus, and EconLit. The search algorithm included all keywords from two groups: (1) "tax," "taxes," "taxation," "taxed," "taxing," "pre-taxation," "posttaxation," "pre-tax," "post-tax," "excise," or "excises"; and (2) "beverage," "beverages," "drink," "drinks," "soda," "sodas," "cola," "coke," "SSB," or "SSBs." The search algorithm used in PubMed was reported in Appendix 1 (Supplementary material). Two co-authors independently screened the title and abstract and identified potentially pertinent articles for the full-text review (Cohen's kappa $\kappa=0.85$ ).

### 2.3. Meta-analysis

Meta-analyses were performed to estimate the pooled effects, represented as mean differences, of soda taxes in the US. The six outcomes included: (1) change in prices of taxed beverages (i.e., SSBs); (2) change in prices of untaxed beverages; (3) change in purchases of taxed beverages; (4) change in purchases of untaxed beverages; (5) change in the consumption of taxed beverages; and (6) tax passthrough rate. Out of the 40 studies, 13 were excluded from the meta-analyses due to: (1) non-overlapping outcome measures ( $36,40,44,56-63$ ), or (2) neither standard error nor confidence interval (CI) reported ( 35,43 ). To assess heterogeneity among the included studies, we employed the $I^{2}$ index, which allows us to quantify the degree of variability between study estimates (64). The $I^{2}$ index was interpreted as modest ( $I^{2} \leq 25 \%$ ), moderate ( $25 \%<I^{2} \leq 50 \%$ ), substantial ( $50 \%<I^{2} \leq 75 \%$ ), or considerable ( $I^{2}>75 \%$ ) (65). Based on the level of heterogeneity observed, we estimated the meta-analyses using either a fixed-effect (FE) model or a random-effect (RE) model. The FE model was utilized when modest or moderate heterogeneity was present, while the RE model was employed when substantial or considerable heterogeneity was observed. To assess publication bias, we conducted Begg's and Egger's tests. These tests allow us to evaluate potential bias in the included studies (66). Additionally, we performed random-effect metaregressions to assess the dose-response effect of alternative soda tax rates on the various outcomes. Meta-analyses were performed using Stata 16.1 MP version. We employed these methods to synthesize the available evidence and provide quantitative estimates of the pooled effects of soda taxes on the specified outcomes.

### 2.4. Study quality assessment

Following Littell et al. (65) and An et al. (67), we designed a study quality assessment tool that rated each study based on ten criteria (Table 5). For each criterion, a score of 1 is assigned if the answer is "yes"; otherwise, a score of 0 is assigned if the answer is "no," "not applicable" or "not reported." We sum the scores of all ten criteria for a total study-specific score, ranging from 0 to 10 . The study quality score was not used as a criterion for inclusion of a study, but as a measure of the strength of the scientific evidence.

## 3. Results

### 3.1. Study selection

Figure 1 shows the study selection flow diagram. We identified 4,574 articles through the keyword and reference search. After removing duplicates, 3,655 articles underwent title and abstract screening, in which 3,562 articles were excluded. The remaining 93 articles were reviewed of full texts against the eligibility criteria. Of these, 53 articles were excluded, including 12 "modeling" studies, 14 "proxy" studies, 17 studies that did not examine soda taxes, six commentaries, and four non-US-based studies. The remaining 40 studies that examined the impact of soda taxes on beverage prices, sales, purchases, and consumption in the US were included in the review (35-52, 56-63, 68-81).

### 3.2. Characteristics of the included studies

Table 1 reports the characteristics of soda taxes assessed in the included studies. All 40 studies were conducted in or after 2015. Fourteen examined soda tax in Philadelphia, Pennsylvania ( $35,39,40$, $43,49,50,52,57,58,60,68,69,76,81$ ); eight in Berkeley, California ( $36,37,44-47,51,56$ ); six in Oakland, California ( $42,61,72,77,78$, 80); three each in Cook County, Illinois (73-75), and Seattle, Washington ( $48,71,79$ ); two in Boulder, Colorado $(38,59)$; one each in two cities-Oakland and San Francisco (70), Berkeley and Washington State (62); one in six cities-Philadelphia, Berkeley, Seattle, Boulder, Cook County, and Oakland (63); and the remaining one in four cities-Philadelphia, San Francisco, Seattle, and Oakland (41). The tax rate ranged from 1 to $2 \phi / o z$. , with a mean of $1.321 \phi / o z$. among the seven cities under investigation.

To account for cross-border estimation across studies within the same jurisdiction, we employed a variety of methods to explain the variation in findings. These methods included considering differences in sample size, sample characteristics, statistical approaches, and outcome measures. By examining these factors, we aimed to provide insights into the varying effects observed within jurisdictions. Table 2 reports sample characteristics, statistical approach, and outcome measures of the studies included in the review. Twenty-nine studies collected beverage pricing or sales data from various types of retailers (e.g., supermarkets, corner/convenience stores, restaurants, and grocery stores) (35-39, 43, 45, 48-51, 56, 59, 61-63, 68-75, 77-81), whereas the remaining 11 surveyed participants regarding their beverage purchases or consumption ( $40-42,44,46,47,52,57,58,60$, 76). All studies except one (73) adopted a difference-in-differences (DID) approach to estimate the impacts of soda taxes. DID is a quasiexperimental approach (82), which uses geographical and timing variations in soda tax implementation across US cities to estimate the causal impact of soda taxes on SSB prices, sales, purchases, and consumption. Twenty-eight studies assessed the effect of soda taxes on SSB or untaxed beverage prices (36-39, 42, 43, 45, 48-51, 59, 6163, 68-73, 75-81). Twenty-two studies focused on beverage sales or purchases ( $35-37,41,42,44,48-51,56,57,62,63,69,73,74,76,77$, 79-81). Nine studies examined the frequency or quantity of beverage consumption ( $40,42,44,46,47,51,52,58,60$ ). Data on beverage prices were collected mainly through three channels: hand-recoding of price tags during retailer visits ( $37-39,42,43,45,59,61,68-72,76$, $78)$, web-scraped data of beverage prices $(38,59)$, or point-of-sale electronic scanner data ( $36,48-51,59,62,63,73,75,77,79-81$ ). Data on beverage sales were collected through two channels: retailers' aggregate sales records $(35,42,56)$ or point-of-sale electronic scanner data ( $36,48-51,62,63,73,74,77,79-81$ ). Data on beverage purchases were collected from surveyed participants ( $40,41,44,57,69,76$ ). Data on beverage consumption were collected through two methods: interviews using a food frequency questionnaire $(40,46,60)$ or a 24 -h dietary recall $(42,51,58)$.

### 3.3. Study findings

Table 3 reports the estimated effects of soda taxes. Four key findings have emerged. First, soda taxes led to increased SSB prices in the taxing jurisdictions relative to the nontaxing jurisdictions. In Berkeley, a $1.0 \not \subset / \mathrm{oz}$. soda tax increased the price for taxed beverages


FIGURE 1
Study selection flow diagram (PRISMA)
by $0.65 ¢ /$ oz. relative to non-Berkeley stores (51). Across all brands and sizes of SSBs, the overall price of taxed beverages increased by 0.43 d/ oz. relative to neighboring cities (37). The price increase was $0.47 \varnothing /$ oz. for small-size SSBs $(\leq 33.8 \mathrm{oz}), 0.46 \not \subset / \mathrm{oz}$. for a 2 -liter SSB bottle, and $0.49 ¢ / \mathrm{oz}$. for a multipack of soda relative to that in neighboring cities (45). In other taxing jurisdictions implementing a $1.0 ¢ / \mathrm{oz}$. soda tax, Falbe et al. documented that the average price of SSBs increased by 0.92 ¢/oz. $(95 \% \mathrm{CI}=0.28,1.56 ; p<0.01)$ in Oakland 10 months post-tax and $1 申 /$ oz. $(95 \% \mathrm{CI}=0.35,1.65 ; \mathrm{p}<0.01)$ in San Francisco 4 months post-tax in comparison to Richmond and San Jose (70) Studies reported that the taxed beverage prices increased by $0.49 \mathrm{~d} / \mathrm{oz}$. (77), and the regular soda price increased by $0.82 \mathrm{c} / \mathrm{oz}$. (72) in Oakland relative to Sacramento one year post-tax. The taxed beverage prices increased by $0.67 ¢ / \mathrm{oz}$. (80), and bottled regular soda prices increased by $1.44 ¢ /$ oz. (78) in Oakland relative to Sacramento two year post-tax. Powell and Leider found that the taxed beverage prices
increased by 1.13 ¢/oz. in Cook County relative to St. Louis (73). Powell et al. found the prices of all taxed beverage types increased by 1.19 ¢/oz. in Cook County relative to St. Louis (75). In Philadelphia, a $1.5 ¢ /$ oz. soda tax raised prices by $1.58 ¢ /$ oz. for all taxed beverages combined relative to comparison untaxed stores outside Philadelphia (39). Coary et al. found that SSB prices increased by $1.53 \mathrm{c} / \mathrm{oz}$. relative to those outside the city (43). Seiler et al. reported that the soda tax passed through at a rate of $97 \%$, leading to a price increase by $1.45 \phi /$ oz. relative to the surrounding area of Philadelphia (50). Bleich et al. reported that the relative price of taxed beverages increased by 1.81 ¢/ oz. 1-year post-tax (69) and 2.06 ¢/oz. two-year post-tax (76) in Philadelphia compared to Baltimore, resulting in a tax passthrough rate of $120.4 \%$ (69) and $137 \%$ (76), respectively. Petimar et al. reported that taxed beverage prices increased by $1.02 \nless / \mathrm{oz}$., with $68 \%$ of the tax passed through to prices 2 years after tax implementation in Philadelphia compared to Baltimore (81). Cawley et al. reported that

TABLE 1 Characteristics of soda taxes assessed in the studies included in the review.

| Study ID | Author, Year | Taxing jurisdiction | Implementation date | Tax rate | Assessment Period |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | Falbe et al. (45) | Berkeley, CA | Jan 1, 2015 | 1 ¢/oz | Pre-tax: Nov 2014 through Jan 2015 <br> Post-tax: May through Jun 2015 |
| 2 | Falbe et al. (46) | Berkeley, CA | Jan 1, 2015 | 1 ¢/oz | Pre-tax: Apr through Jul 2014 <br> Post-tax: Apr through Aug 2015 |
| 3 | Cawley and Frisvold (37) | Berkeley, CA | Jan 1, 2015 | 1 ¢/oz | Pre-tax: Dec 22, 2014 Post-tax: Jun 1, 2015 |
| 4 | Debnam (44) | Berkeley, CA | Jan 1, 2015 | 1 / $/ \mathrm{oz}$ | 2010 through 2015 |
| 5 | Silver et al. (51) | Berkeley, CA | Jan 1, 2015 | 1 / $/ \mathrm{oz}$ | 1. Prices data: <br> Pre-tax: Dec 2014 <br> Post-tax: Jun 2015 and Mar 2016 <br> 2. Point of electronic sale data: <br> Jan 1, 2013, through Feb 29, 2016 |
| 6 | Bollinger and Sexton (36) | Berkeley, CA | Jan 1, 2015 | 1 / $/ \mathrm{oz}$ | Jan 2013 through Dec 2015 |
| 7 | Cawley et al. (68) | Philadelphia, PA | Jan 1, 2017 | 1.5 ¢/oz | Pre-tax: December 21, 2016 <br> Post-tax: January 14, 2017, and February 5, 2017 |
| 8 | Cawley et al. (68) | Boulder, CO | Jan 1, 2017 | 2 ¢/oz | Pre-tax: Apr through Jun 2017 <br> Post-tax: Aug through Oct 2017 |
| 9 | Coary and Baskin (43) | Philadelphia, PA | Jan 1, 2017 | 1.5 ¢/oz | Pre-tax: Dec 2016 <br> Post-tax: Sep 2017 |
| 10 | Zhong et al. (52) | Philadelphia, PA | Jan 1, 2017 | 1.5 ¢/oz | Pre-tax: Dec 6 through Dec 312016 <br> Post-tax: Jan 15 through Feb 31, 2017 |
| 11 | Baskin and Coary (35) | Philadelphia, PA | Jan 1, 2017 | 1.5 ¢/oz | Pre-tax: Nov 2015 through Feb 2016 Post-tax: Nov 2016 through Feb 2017 |
| 12 | Cawley et al. (41) | 1. Philadelphia, PA <br> 2. San Francisco, CA <br> 3. Seattle, WA <br> 4. Oakland, CA | 1. Jan 1, 2017 <br> 2. Jan 1,2018 <br> 3. Jan 1,2018 <br> 4. Jul 1, 2017 | $\begin{gathered} 1.1 .5 \phi / o z . \\ 2.1 \phi / \mathrm{oz} \text {. } \\ \text { 3. } 1.75 \phi / \mathrm{oz} . \\ 4.1 \not / \mathrm{oz} \end{gathered}$ | Jul 1, 2016 through Jun 30, 2017 |
| 13 | Cawley et al. (40) | Philadelphia, PA | Jan 1, 2017 | 1.5 ¢/oz | Pre-tax: Nov through Dec 2016 <br> Post-tax: Nov through Dec 2017 |
| 14 | Lee et al. (47) | Berkeley, CA | Jan 1, 2015 | 1 / $/ \mathrm{oz}$ | Pre-tax: Apr through Jul 2014 <br> Post-tax: Apr through Oct 2017 |
| 15 | Roberto et al. (49) | Philadelphia, PA | Jan 1, 2017 | 1.5 ¢/oz | Jan 1, 2014 through Dec 31, 2017 |
| 16 | Taylor et al. (56) | Berkeley, CA | March 1, 2015 (not on campus) <br> August 2016 (on campus) | 1 / $/ \mathrm{oz}$ | The pre-election campaign period: July 2014 through October 2014 <br> The postelection and pretax implementation period: November 2014 through February 2015 <br> The tax implementation period not on campus: March 2015 through July 2016 <br> The tax implementation period on campus: August 2016 through December 2016 |
| 17 | Bleich et al. (69) | Philadelphia, PA | Jan 1, 2017 | 1.5 ¢/oz | Pre-tax: Oct through Dec 2016 <br> Post-tax: Jun through Aug 2017, Oct through Dec 2017 |
| 18 | Cawley et al. (40) | Oakland, CA | Jul 1, 2017 | 1 ¢/oz | Pre-tax: Apr through Jun 2017 <br> Post-tax: Apr through Jun 2018 |
| 19 | Cawley et al. (40) | Philadelphia, PA | Jan 1, 2017 | 1.5 ¢/oz | Pre-tax: Nov through Dec 2016 <br> Post-tax: Nov through Dec 2017 |

(Continued)

TABLE 1 (Continued)

| Study ID | Author, Year | Taxing jurisdiction | Implementation date | Tax rate | Assessment Period |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 20 | Falbe et al. (70) | Oakland, CA <br> San Francisco, CA | Jul 1, 2017 <br> Jan 1, 2018 | 1 ¢/oz | Pre-tax: Apr through May 2017 <br> Post-tax: Apr through May 2018 |
| 21 | Jones-Smith et al. (71) | Seattle, WA | Jan 1, 2018 | 1.75 ¢/oz | Pre-tax: Oct through Nov 2017 <br> Post-tax: May through Jul 2018 |
| 22 | Lawman (57) | Philadelphia, PA | Jan 1, 2017 | 1.5 ¢/oz | Pre-tax: Sept through Dec 2016 <br> Post-tax: 3,6 , and 12 months after tax implementation |
| 23 | Marinello et al. (78) | Oakland, CA | Jul 1, 2017 | 1 / oz | Pre-tax: May through Jun 2017 <br> Post-tax: Jan 2018 and Jun 2018 |
| 24 | Powell and Leider (73) | Cook County, IL | August 2, 2017 | 1 ¢/oz | Pre-tax: March 29, 2015 through July 29, 2017 <br> Post-tax: August 6, 2017 through November 25, 2017 <br> Post-repeal: December 3, 2017 through <br> August 4, 2018 |
| 25 | Powell et al. (74) | Cook County, IL | Aug 2, 2017 | 1 ¢/oz | Pre-tax: Aug 7, 2016, to Nov 26, 2016 <br> Post-tax: Aug 6, 2017, to Nov 25, 2017 |
| 26 | Powell and Leider (48) | Seattle, WA | Jan 1, 2018 | 1.75 ¢/oz | Pre-tax: Sep 29, 2017 through Feb 4, 2018 Post-tax: Feb 4, 2018 through Sep 29, 2018 |
| 27 | Powell et al. (75) | Cook County, IL | Aug 2, 2017 | 1 ¢/oz | Pre-tax: Aug 7, 2016, to Nov 26, 2016 <br> Post-tax: Aug 6, 2017, to Nov 25, 2017 |
| 28 | Zhong et al. (58) | Philadelphia, PA | Jan 1, 2017 | 1.5 ¢/oz | Pre-tax: Dec 2016 through Jan 2017 <br> Post-tax: Dec 2017 through Feb 2018 |
| 29 | Bleich et al. (58) | Philadelphia, PA | Jan 1, 2017 | 1.5 ¢/oz | Baseline: October through December 2016 Post-tax: 6, 12, and 24 months after tax implementation |
| 30 | Cawley et al. (59) | Boulder, CO | July 1, 2017 | 2 ¢/oz | Pre-tax: April and June 2017 <br> Post-tax: August and October 2017 |
| 31 | Edmondson et al. (60) | Philadelphia, PA | Jan 1, 2017 | $1.5 ¢ / \mathrm{oz}$ | September 2012 through December 2019 |
| 32 | Léger and Powell (77) | Oakland, CA | Jul 1, 2017 | 1 ¢/oz | Pre-tax: July 1, 2016 to June 30, 2017 <br> Post-tax: July 1, 2017 to June 30, 2018 |
| 33 | Léger and Powell (77) | Oakland, CA | Jul 1, 2017 | 1 ¢/oz | Pre-tax: late May-June 2017 <br> Post-tax: June 2019 |
| 34 | Marinello et al. (78) | Oakland, CA | Jul 1, 2017 | 1 ¢/oz | Pre-tax: May through June 2017 <br> Post-tax: June 2019 |
| 35 | Powell et al. (53, 54) | Seattle, WA | Jan 1, 2018 | 1.75 ¢/oz | Pre-tax: 2017 <br> Post-tax: February 3, 2019 through <br> September 28, 2019 |
| 36 | Rojas and Wang (62) | 1. Berkeley, CA <br> 2. Washington State | 1. March 1, 2015 <br> 2. July 1, 2010 | $\begin{gathered} 1 \text { ¢/oz. } \\ 0.166 \text { d/oz } \end{gathered}$ | 1. January 2014 through December 2015 <br> 2. January 2009 through December 2012 |
| 37 | Seiler et al. (50) | Philadelphia, PA | Jan 1, 2017 | 1.5 ¢/oz | Jan 2015 through Sep 2018 |
| 38 | Zhang et al. (63) | 1. Seattle, WA <br> 2. Boulder, CO <br> 3. Cook County, IL <br> 4. Philadelphia, PA <br> 5. Berkeley, CA <br> 6. Oakland, CA | 1. Jan 2018 <br> 2. Jul 2017 <br> 3. Aug 2017 <br> 4. Jan 2017 <br> 5. Mar 2015 <br> 6. Jul 2017 |  | From 2013 to 2018 |
| 39 | Leider and Powell (80) | Oakland, CA | Jul 1, 2017 | 1 ¢/oz | Pre-tax: July 31, 2016 through May 27, 2017 <br> Post-tax: July 29, 2018 through May 25, 2019 |
| 40 | Petimar et al. (81) | Philadelphia, PA | Jan 1, 2017 | $1.5 ¢ / \mathrm{oz}$ | From January 1, 2016 to December 30, 2018 |

TABLE 2 Sample characteristics, sample size, and measures of the studies included in the review.

| Study ID | Sample size | Sample characteristics | Statistical approach | Outcomes assessed | Outcome measures |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 71 retailers | Chain supermarket, small grocery store, drugstore, convenience store, liquor store | DID | Changes in prices of SSBs and non-SSBs | 1. Trained research assistants collected beverage prices by recording visible prices from the price tag <br> 2. For beverages without visible prices, data collectors asked store clerks for prices <br> 3. If clerks were uncooperative, data collectors purchased beverages and recorded prices from receipts <br> 4. If a temporary promotional price was advertised, data collectors recorded both the promotional and regular price |
| 2 | 2,679 participants | Low-income and minority population | DID | Changes in beverage consumption | Beverage consumption via interviewer-administered intercept surveys with a beverage frequency questionnaire modified from the Behavioral Risk Factor Surveillance System 2011 SSB module |
| 3 | 86 stores | Supermarkets, grocery stores, pharmacies, convenience stores, and gas stations with posted prices | DID | Price of SSBs | Store visits |
| 4 | 2,399,897 household-purchase-weeks | Households | DID (fixed- <br> effect <br> regression) | Household purchases | 1. Data from the Nielsen <br> Consumer Panel <br> 2. For purchases made at retailers where Nielsen does not receive the point of sale data, the panelist manually enters the expenditure made on the purchase |
| 5 | 1. 26 stores <br> 2. 2,175 store prices <br> 3. Sales data covered 118.8 million barcode scans from 15.5 million transactions | 1. Large supermarkets, small chain supermarkets, chain and independent gas stations, pharmacies, and independent corner stores <br> 2. Store price collected for a standard panel of 70 beverages, which included 45 taxed and untaxed branded beverages <br> 3. Point-of-sale electronic scanner data | DID (fixed- <br> effect <br> regression) | Changes in prices, <br> beverage sales, and usual beverage intake | 1. Data collection protocols were employed to measure beverage prices systematically <br> 2. Point-of-sale electronic scanner data were requested using personal outreach <br> 3. Trained interviewers used standardized questionnaires and computer-assisted telephone interviews to conduct a 24 -h recall of beverage intake <br> 4. A second 24 -h beverage recall interview was collected 3-7 days later |
| 6 | 1. 3,549 UPCs <br> 2. 196,226 weekly UPC <br> prices and sale quantities | UPCs in retail chains | DID | Weekly average prices, <br> sales quantities and UPC <br> (product) volume | Nielsen scanner data |
| 7 | 31 stores | Retail chains (including bakeries, restaurants, and newsstands) | DID | Beverage prices | Data collected in person at retail stores |

(Continued)

TABLE 2 (Continued)

| Study ID | Sample size | Sample characteristics | Statistical approach | Outcomes assessed | Outcome measures |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 8 | 1. Retailers: 1,035 <br> 2. Restaurants: 1,263 <br> 3. OrderUp: 158 | All retail stores, all limited-service restaurants, and a selected sample of restaurant menus | DID | The changes in beverage prices | 1. Hand-collected data of listed prices and purchase prices of beverages from all retail stores <br> 2. Hand-collected data of listed prices of fountain drinks and coffee drinks from all limitedservice restaurants <br> 3. Web-scraped data of prices from a selected sample of restaurant menus |
| 9 | 249 products: 190 soda products, 38 juice products, and 21 water products | Soda, juice, and water | DID | Beverage prices | Data collected from 2 matched stores from each category both inside and outside Philadelphia county |
| 10 | 1,777 participants | Residents | DID | The daily quantity of consumption, and 30 -day average consumption frequency and quantity | Phone-based survey: Beverage questions were based on a modified version of the 15-item Beverage Intake Questionnaire |
| 11 | 9 stores | Predominant grocery retailers <br> (top 3 in total retail sales) | DID | Total beverage sales by category | Data came from the retailers |
| 12 | 1,447 households | Households with children | DID | Households' monthly beverage purchases | Data collected by InfoScout |
| 13 | 1. Purchase: 1,305 individuals in 2016 and 1,501 in 2017 <br> 2. Household survey: 440 | Consumers at stores, households with children, and residents | DID | 1. Changes in the volume of taxed and untaxed beverages purchased <br> 2. Beverage consumption | 1. Interviewed consumers at stores <br> 2. Household survey of beverage consumption based on telephone and online using the National Health and Nutrition Examination Survey Dietary Screener Questionnaire |
| 14 | 5,225 participants | Residents | DID | SSB and water consumption | SSB consumption was measured annually through beverage frequency questionnaires. The BFQs were based on the previously validated BEVQ-15 |
| 15 | 291 stores | Chain retailers: supermarket, mass merchandizers, and pharmacies | DID | Change in taxed beverage prices and volume sales | Data were purchased from IRI |
| 16 |  | University retailers | DID | The change in soda sales | Panel data of beverage sales from university retailers |
| 17 | 134 stores and 4,584 purchases | Independent stores | DID | Beverage prices and purchases | Research staff documented beverage prices and collected purchase data from consumers |
| 18 | 126 stores | Stand-alone convenience stores, gas stations with convenience stores, small and large grocery stores, and pharmacies | DID | Beverage prices, purchases, and consumption | Data collected from stores |
| 19 | 1. 3,152 prices for 2016 <br> 2. 2,763 prices for 2017 | Prices from both taxed beverage and untaxed beverage | DID | Changes in beverage prices | Data collected in person at retail stores |

(Continued)

TABLE 2 (Continued)

| Study ID | Sample size | Sample characteristics | Statistical approach | Outcomes assessed | Outcome measures |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 20 | 155 stores | Supermarkets, drugstores, mass merchandizers, convenience stores, corner and small grocery stores, and liquor stores | DID (fixed- <br> effect <br> regression) | Retail prices of SSBs and non-SSBs | Data collectors recorded beverage shelf prices in stores |
| 21 | 407 stores | Retail food stores, quick-service restaurants, and coffee shops | DID | Beverage prices | Data collected from stores and restaurants |
| 22 | 603 participants | Adult SSB consumers | DID (mixed- <br> effect <br> regression) | Beverage purchases | Receipts collected from consumers |
| 23 | 150 observations of bottled regular soda from 39 restaurants, 106 observations of bottled diet soda from 32 restaurants, and 501 observations of fountain drinks from 73 restaurants | Fast-food restaurants | DID | Beverage prices | Data were collected in-person using the Beverage Tax Fast-food Restaurant Observation Form |
| 24 | 16, 510 UPCs for volume and 2,141 UPCs for price | Taxed and untaxed UPCs | Interrupted time series | Changes in beverage prices and volume sold | Nielsen food store scanner data |
| 25 | 7,798 UPCs | Taxed and untaxed UPCs within the sites and in the 2-mile border area of each site | DID | The volume sold of taxed and untaxed beverages | Retail scanner data on sales of non-alcoholic beverages were obtained from the Nielsen panel survey |
| 26 | 7,868 UPCs | Taxed and untaxed UPCs within the sites and in their 2-mile border areas | DID | Beverage prices and volume sold | UPC-level retail store scanner data obtained from the Nielsen |
| 27 | 13,015 UPCs | Distinct beverage UPCs sold | DID | Beverage prices | Retail scanner data on sales of nonalcoholic beverages were obtained from the Nielsen panel survey |
| 28 | 515 participants | Residents aged 18-64 years | DID | Change in beverage consumption | A random-digit-dialing phone survey was administered to a population-based cohort |
| 29 | 116 independent stores and 4,738 customer purchases | Independent stores | DID | Changes in the prices and purchases of beverages | Data were collected by research assistants |
| 30 | 621 retailers <br> 36 stores <br> 776 restaurants <br> 219 restaurants | Retailers, liquor stores, restaurants | DID | Pass-through of sugar sweetened beverages taxes | Data were collected from four sources: (a) handcollected data on prices from stores; (b) Nielsen Retail Scanner Data of store prices; (c) hand-collected data on prices in restaurants; and (d) web-scraped data from online restaurant menus. |
| 31 | 86,928 participants | High School Students | DID | Soda consumption | Data were obtained from Youth <br> Risk Behavior Surveillance System |
| 32 | 2,187 UPCs | Taxed and untaxed UPCs in supermarkets, grocery stores, convenience stores, drug stores, mass merchandize stores, and dollar stores | DID | Changes in beverage prices and volume sold | Data were obtained from customordered Nielsen retail scanner data |

TABLE 2 (Continued)

| Study ID | Sample size | Sample characteristics | Statistical approach | Outcomes assessed | Outcome measures |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 33 | 210 stores | Seven store types: general merchandize stores, supermarkets, grocery stores, chain convenience stores, non-chain convenience stores, small discount stores, and drug stores/pharmacies | DID | Change in beverage prices | Data were collected from stores by the data collector |
| 34 | 85 restaurants | Chain and non-chain fast-food restaurants | DID | Change in prices of beverages sold | Data were collected at restaurants |
| 35 | 6,652 UPCs | Taxed and untaxed UPCs in Grocery, drug, convenience, dollar, and mass merchandize stores, and supermarkets | DID | Changes in beverage prices and volume sold | Data were obtained from Nielsen |
| 36 | 1,503 stores | Grocery stores, drug stores, mass merchandizers, etc. | DID | Changes in beverage prices and volume sold | Data were obtained from the Nielsen retail scanner database |
| 37 | 357 stores | Chain stores | DID | Beverage prices and quantities sold | Retail scanner data collected by IRI |
| 38 | 3,719 stores | Supermarkets and convenient stores | DID | Changes in beverage prices and sales quantity | Data from IRI, the InfoScan, store-based scanner data |
| 39 | 11,705 UPCs | Distinct beverage UPCs sold | DID | Changes in beverage prices and volume sold | Data were obtained from Nielsen |
| 40 | 109 supermarkets, 45 mass merchandizers, and 350 pharmacies | Large retailers | DID | Changes in beverage price and sales | Data from IRI |

IRI, Information Resources, Inc.; DID, difference-in-differences; SSB, sugar-sweetened beverage.
the price of taxed beverages increased by $0.83 ¢ /$ oz. relative to untaxed beverages, and $55.3 \%$ of the tax was passed through to consumers at the Philadelphia International Airport (68). In Seattle, a $1.75 \phi /$ oz. soda tax raised prices of taxed beverages by $1.03 \phi /$ oz. relative to Portland, corresponding to a $59 \%$ tax passthrough rate (48). JonesSmith et al. found that the soda tax was associated with a price increase of $1.58 \phi / \mathrm{oz}$. relative to its comparison area south of Seattle, and $90 \%$ of the tax was passed through to consumers (71).

Second, soda taxes reduced SSB sales and purchases in the taxing jurisdictions relative to the nontaxing jurisdictions. SSB sales in Berkeley declined by $2.7 \%$ relative to neighboring cities (51). The total sales volume of taxed beverages decreased by $48.7 \%$ (49), or $38.9 \%$ (69) in Philadelphia relative to Baltimore following soda tax implementation. Similarly, the volume sold of taxed beverages decreased by $27 \%$ ( 74 ) or $25.7 \%$ (73) in Cook County relative to St. Louis. Sales volume of taxed SSBs reduced by $14 \%$ (77) or $18 \%$ (80) in Oakland compared to Sacramento. Taxed beverage volume sales decreased by $50 \%$ in Philadelphia relative to Baltimore two years post-tax (81). Taxed beverage sales fell by $22 \%$ in Seattle relative to Portland (79). Compared with the surrounding metropolitan area, sales quantity of taxed SSB fell by $10.6 \%$ in Berkeley, $14.4 \%$ in Oakland, $5.1 \%$ in Boulder, $8.3 \%$ in Philadelphia, $10.7 \%$ in Cook County, and $5.6 \%$ in Seattle, respectively (63). Concerning purchases, taxed beverage purchases declined by 6.1 floz ., corresponding to a $42 \%$ decline in Philadelphia relative to Baltimore (76). An increase in the soda tax rate by $1 \phi /$ oz. decreased monthly household purchases of taxed beverages by 53.0 oz . or $12.2 \%$ (41). The average quantity of
taxed beverages purchased per shopping trip decreased in Philadelphia and increased in neighboring cities, which resulted in a relative decrease of 8.5 oz . of SSBs purchased per shopping trip in Philadelphia (40).

Third, soda taxes reduced the consumption of SSBs in taxing jurisdictions relative to nontaxing jurisdictions. Consumption of SSBs was found to decrease by $25 \%$ in Berkeley relative to neighboring cities (46). Soda intake was reduced by 0.81 servings per week in Philadelphia compared with all other comparison cities two years after tax implementation (60). The odds of daily consumption of regular soda and energy drinks were 40 and $64 \%$ lower, respectively, in Philadelphia relative to the comparison cities two-month post-tax (52). Adults in Philadelphia were found to consume regular soda 10.4 fewer times per month after soda tax implementation, denoting a reduction of approximately $30 \%$ (40). People in Philadelphia were more likely to reduce their frequency of SSB consumption, and their monthly SSB consumption declined by 51.65 oz . compared to those residing in neighboring cities (58).

Fourth, the impact of soda taxes on SSB sales and purchases was compromised by cross-border shopping. Roberto et al. estimated that approximately a quarter of the decrease in the taxed beverage sales in Philadelphia could be offset by increases in sales in bordering areas (49). Seiler et al. estimated that cross-shopping to stores outside Philadelphia offset over half of the reduction in SSB sales in the city (50). Petimar et al. estimated that $30 \%$ of the reduction in taxed beverage sales in Philadelphia was offset by cross-border shopping

| Study ID | Effects of a soda tax on consumption | Effects of a soda tax on sales or purchases | Effects of a soda tax on prices |
| :---: | :---: | :---: | :---: |
| 1 |  |  | 1. For small-size beverages ( $\leq 33.8 \mathrm{oz}$ ), price increases in Berkeley relative to those in comparison cities were $0.69 ~ ¢ / \mathrm{oz}$. $(95 \% \mathrm{CI}=0.36,1.03)$ for soda, $0.47 \phi /$ oz. $(95 \% \mathrm{CI}=0.08,0.87)$ for fruit-flavored beverages, and $0.47 \not \subset / \mathrm{oz}$. $(95 \%$ $\mathrm{CI}=0.25,0.69$ ) for SSBs overall. A pass-through rate of $47 \%$ for SSBs overall. <br> 2. For 2-liter bottles and multipacks of soda, relative price increases were $0.46 \not \subset /$ oz. ( $95 \% \mathrm{CI}=0.03,0.89$ ) and 0.49 ф/oz. ( $95 \% \mathrm{CI}=0.21,0.77$ ). <br> 3. No relative price increases for untaxed beverages overall. |
| 2 | 1. Consumption of SSBs decreased by $21 \%$ in Berkeley and increased by $4 \%$ in comparison cities ( $p=0.046$ ), the ratio of post- to pre-tax consumption in Berkeley relative to comparison cities is $0.76,95 \% \mathrm{CI}=0.58,0.995$. <br> 2. Water consumption increased more in Berkeley ( $+63 \%$ ) than in comparison cities ( $+19 \%$; p $<0.01$ ). |  |  |
| 3 |  |  | 1. Across all brands and sizes of products examined, $43.1 \%$ ( $95 \% \mathrm{CI}=27.7$, $58.4 \%$ ) of the Berkeley tax was passed on to consumers. <br> 2. For each mile of distance between the store and the closest store selling untaxed SSBs, pass-through rose $33.3 \%$ for 2-liter bottles and $25.8 \%$ for 12 -packs of $12-\mathrm{oz}$ cans. |
| 4 | Households living in Alameda County increased their consumption of sugar-sweetened beverages by 8.89 oz . relative to other households in the U.S., increased their consumption of soda by 26.56 oz . relative to other households in the U.S., and by 37.15 oz. relative to residents of San Mateo, a neighboring county. |  |  |
| 5 | Reductions in self-reported mean daily SSB intake in grams $(-19.8 \%, p=0.49)$ and mean per capita SSB caloric intake ( $-13.3 \%, p=0.56$ ) from baseline to post-tax were not statistically significant. | 1. SSB sales in Berkeley stores declined 9.6\% ( $\mathrm{p}<0.001,95 \% \mathrm{CI}=-9.9$, $-9.3 \%)$ compared to estimates if the tax were not in place, but rose $6.9 \%$ ( $p<0.001,95 \% \mathrm{CI}=6.3,7.2$ ) for non-Berkeley stores. <br> 2. Sales of untaxed beverages in Berkeley stores rose by $3.5 \% ~(95 \% \mathrm{CI}=3.1$, 3.9) versus $0.5 \%(95 \% \mathrm{CI}=0.1,0.9)$ (both $p<0.001$ ) for non-Berkeley stores. | 1. Pass-through was complete in large chain supermarkets ( +1.07 ¢/oz., $\mathrm{p}=0.001$ ) and small chain supermarkets and chain gas stations ( $1.31 申 / \mathrm{oz}$., $\mathrm{p}=0.004$ ), partial in pharmacies ( $+0.45 \phi / \mathrm{oz} ., p=0.03$ ), and negative in independent corner stores and independent gas stations ( $-0.64 \phi / \mathrm{oz},. p=0.004$ ). <br> 2. Sales-unweighted mean price change from scanner data was $+0.67 \mathrm{q} / \mathrm{oz}$. ( $p<0.001$ ) (sales-weighted, $+0.65 \phi /$ oz., $p=0.003,95 \% \mathrm{CI}=0.23,1.07$ ), with +1.09 //oz. ( $\mathrm{p}<0.001$ ) for sodas and energy drinks, but a lower change in other categories. <br> 3. One year following implementation of the nation's first large SSB tax, prices of SSBs increased in many, but not all, settings. |
| 6 |  | Reduced supermarket purchases of soda in the taxing jurisdiction. Half of these reduced purchases are substituted to just outside the taxing jurisdiction. | Relative to an in-state synthetic control, per-ounce prices and volume-weighted per-ounce prices at the Berkeley supermarket are estimated to increase by a statistically significant $0.19 \phi$ and $0.15 \phi$, respectively. These constitute price increases of 2.95 and $4.35 \%$, respectively, reflecting respective tax pass-through rates of 18.53 and $15.25 \%$. |

TABLE 3 (Continued)

| Study ID | Effects of a soda tax on consumption | Effects of a soda tax on sales or purcha | fects of a soda tax on pric |
| :---: | :---: | :---: | :---: |
| 7 |  |  | One month after tax implementation, prices had increased by $0.83 \phi / \mathrm{oz}$. $(95 \%$ $\mathrm{CI}=0.33,1.33 ; \mathrm{p}=0.002$ ) in taxed stores relative to untaxed stores, and $55.3 \%$ ( $95 \% \mathrm{CI}=22,89 \% ; p=0.002$ ) of the tax was passed on to consumers. |
| 8 |  |  | 1. The tax was substantially, but not entirely, passed through to consumers in the form of higher prices; pass-through rates can vary across different localities. <br> 2. The estimated pass-through based on posted prices at stores is $51.2 \%$; whereas, pass-through based on register prices is $79.3 \%$. <br> 3. Data hand-collected from restaurants indicates that the pass-through of the tax was $69.4 \%$ on fountain drinks. <br> 4. There is little evidence of any impact of the tax on the store prices of untaxed beverages. <br> 5. The change of the taxed products retail register price relative to the prices in June in Boulder County and Fort Collins is $1.550, \mathrm{SE}=0.201$. |
| 9 |  |  | Outside Philadelphia County, soda products prices increased, on average $\$ 0.0012$ / oz., while prices inside Philadelphia County prices increased, on average, $\$ 0.0165$. The difference between these two increases is $\$ 0.0153 .100 \%$ excise tax passthrough rate on SSBs with almost no price change on substitute products. |
| 10 | Within the first 2 months of tax implementation, relative to the comparison cities, in Philadelphia the odds of daily consumption of regular soda was $40 \%$ lower ( $\mathrm{OR}=0.6,95 \% \mathrm{CI}=0.37,0.97$ ); energy drink was $64 \%$ lower ( $\mathrm{OR}=0.36,95 \% \mathrm{CI}=0.17,0.76$ ); bottled water was $58 \%$ higher $(\mathrm{OR}=1.58,95 \% \mathrm{CI}=1.13,2.20)$; and the 30 -day regular soda consumption frequency was $38 \%$ lower ( $\mathrm{OR}=0.62,95 \% \mathrm{CI}=0.40,0.98$ ). |  |  |
| 11 |  | 1. There is a significant difference in the change of total beverage sales for those inside Philadelphia County versus those outside Philadelphia County ( $\mathrm{t}=5.35, \mathrm{p}<0.001$ ). <br> 2. Stores inside Philadelphia County experienced a significantly higher change (decrease in total beverage sales) in sales than those outsides of Philadelphia County (increase in total beverage sales). <br> 3. This change did not coincide with a corresponding change in untaxed beverages. |  |
| 12 |  | 1. An increase in the beverage tax rate of $1 \phi / \mathrm{oz}$. decreases household purchases of taxed beverages by 53.0 oz . per month or $12.2 \%$. <br> 2. The decline concentrated in Philadelphia, where the tax decreased purchases by $27.7 \%$, and there was no impact of the taxes in the other three cities combined. |  |


| Study ID | Effects of a soda tax on consumption | Effects of a soda tax on sales or purchases | Effects of a soda tax on prices |
| :---: | :---: | :---: | :---: |
| 13 | 1. The estimated impact of a soda tax on the consumption of added sugars from SSBs and the frequency of consuming taxed beverages are negative but not statistically significant for children (a reduction of $2.4 \mathrm{~g} /$ day) and adults (a reduction of $5.9 \mathrm{~g} /$ day). <br> 2. Due to soda tax, adults in Philadelphia consumed regular soda 10.4 fewer times per month, which is a reduction of approximately $30 \%$. | The average amount of taxed beverages purchased per shopping trip decreased in Philadelphia and increased in comparison communities, which results in a relative decrease of 8.5 oz . per shopping trip at stores in Philadelphia. |  |
| 14 | 1. At baseline, SSBs were consumed 1.25 times/day ( $95 \%$ $\mathrm{CI}=1.00,1.50$ ) in Berkeley and 1.27 times/day ( $95 \% \mathrm{CI}=1.13$, <br> 1.42) in comparison city. <br> 2. Adjusting for covariates, consumption in Berkeley declined by 0.55 times/day ( $95 \% \mathrm{CI}=-0.75,-0.35$ ) for SSBs and increased by 1.02 times/day $(95 \% \mathrm{CI}=0.54,1.50)$ for water. <br> 3. Changes in consumption in Berkeley were significantly different from those in the comparison group, which had no significant changes. |  |  |
| 15 |  | 1. In Philadelphia and Baltimore, taxed beverage volume sales in 4 weeks decreased in all store types. <br> 2. Compared with Baltimore, there were significantly more substantial declines in the volume of taxed beverages sold in the after-tax period in Philadelphia. Total volume sales of taxed beverages in Philadelphia decreased by 1.3 billion oz. (from 2.475 billion to 1.214 billion) or by $51.0 \%$ following tax implementation. | 1. In Philadelphia and Baltimore, the mean price per oz. of taxed beverages increased at all stores in the after-tax periods. <br> 2. Compared with Baltimore, Philadelphia experienced significantly greater increases in taxed beverage prices. For supermarkets, the prices increased 0.65 $\$ /$ oz., $95 \% \mathrm{CI}=0.60,0.69$; For mass merchandize stores, the price increased 0.87 $\phi /$ oz., $95 \% \mathrm{CI}=0.72,1.02$; For pharmacies, the price increased 1.56 \$/oz., $95 \%$ $\mathrm{CI}=1.50,1.62$. |
| 16 |  | Soda sales fell significantly compared with control beverage groups in the period immediately following the election, decreasing by between 10 and $20 \%$ compared with precampaign levels. On-campus soda sales continued to fall when the tax was implemented in the city but not on campusdecreasing by $18-36 \%$ compared with the precampaign period-and remained at this depressed level after the tax implementation on campus. |  |
| 17 |  | The volume sold of taxed beverages declined by $38.9 \%$ in Philadelphia after soda tax implementation compared to Baltimore. | The price increase of taxed beverages in Philadelphia was $1.81 申 /$ oz. (95\% $\mathrm{CI}=1.52,2.09$ ) compared to Baltimore, revealing a $120.4 \%$ tax pass-through rate. |
| 18 | No evidence showed that substantial changes in the overall consumption of SSBs or added sugars consumed through beverages for either adults or children after the tax. | There was a slight decrease in the volume of SSBs purchased per shopping trip in Oakland and a small increase in purchases at stores outside of the city, resulting in a decrease in purchases of 11.33 oz ., but it was not statistically significant. | Roughly $60 \%$ of the tax passed on to consumers in the form of higher prices. |


| Study ID | Effects of a soda tax on consumption | Effects of a soda tax on sales or purchases | Effects of a soda tax on prices |
| :---: | :---: | :---: | :---: |
| 19 |  |  | 1. The tax was fully passed through to consumers via higher retail prices. <br> 2. For all taxed beverages combined, the 1.5 \&-per-oz tax raised prices by 1.582 ф/oz. ( $95 \% \mathrm{CI}=1.21,1.89$ ). <br> 3. The impact of the tax is substantial, raising prices per oz. by $21 \%$ on average. Pass-through is complete for specific categories of taxed beverages, such as regular soda ( $1.591 申 / \mathrm{oz}$ ), diet soda ( $1.551 申 / \mathrm{oz}$ ), energy drinks ( $1.998 \phi / \mathrm{oz}$ ), and juice drinks ( $1.928 \phi / o z$ ). |
| 20 |  |  | 1. The average price of SSBs increased by $0.92 \not \subset / \mathrm{oz}$. $(95 \% \mathrm{CI}=0.28,1.56)$ in Oakland and $1.00 \$ / \mathrm{oz}$. ( $95 \% \mathrm{CI}=0.35,1.65$ ) in San Francisco, compared to prices in untaxed cities. <br> 2. The soda tax did not significantly alter prices of water, $100 \%$ juice, or milk of any size examined. <br> 3. Diet soda exhibited a higher price increase in taxed cities. |
| 21 |  |  | 1. The soda tax was associated with an average price increase of $1.58 \phi /$ oz. among Seattle retailers. Nearly the full cost of the tax was passed through to consumers in Seattle. <br> 2. Prices of some non-taxed beverages also increased while the prices of healthy foods generally did not. |
| 22 |  | The soda tax was associated with reductions of taxed beverage purchases at 3 and 6 month but not 12 month. Analyses aggregating all 6 weeks of post-tax time points showed significant reductions ( $-203.7 \mathrm{oz} ., 95 \%$ $\mathrm{CI}=-399.6,-7.8$ ). |  |
| 23 |  |  | 1. There was a $82 \%$ tax pass-through for bottled regular soda one year after tax implementation, with a price increase of $8 \%$. <br> 2. No statistically significant change in prices was found in either time period for taxed and untaxed fountain drinks and untaxed bottled diet soda. |
| 24 |  | Volume sold of taxed beverages in Cook County compared with St Louis exhibited a posttax implementation level decrease of $25.7 \%$ ( $\beta=-0.297$; $95 \% \mathrm{CI}=-0.415,-0.179$ ) and a posttax repeal level increase of $30.5 \%(\beta=0.266,95 \% \mathrm{CI}=0.124,0.408)$, with no net change in volume sold from pretax to 8 months after repeal. | Compared with St Louis, posttax implementation in Cook County resulted in a level increase in taxed beverage prices of $1.13 \phi / \mathrm{oz}$. ( $95 \% \mathrm{CI}=1.01,1.25$ ), followed by a posttax repeal level decrease of $-1.19 ¢ /$ oz. $(95 \% \mathrm{CI}=-1.33$, -1.04 ), with no resulting change pretax to posttax repeal. |


| Study ID | Effects of a soda tax on consumption | Effects of a soda tax on sales or purchases | Effects of a soda tax on prices |
| :---: | :---: | :---: | :---: |
| 25 |  | 1. Volume sold of taxed beverages decreased by $27 \%$ ( $95 \% \mathrm{CI}=-30,-25 \%$ ) in Cook County relative to St. Louis. <br> 2. The magnitude of decrease in volume sold across types of taxed beverages was heterogeneous: $-32 \%(95 \% \mathrm{CI}=-35,-28 \%)$ for soda versus $-11 \%(95 \% \mathrm{CI}=-18,-3 \%)$ for energy drinks, $-37 \%(95 \% \mathrm{CI}=-41,-34 \%)$ for diet beverages versus $-25 \%(95 \% \mathrm{CI}=-28,-21 \%)$ for SSBs , and $-29 \%$ ( $95 \% \mathrm{CI}=-32,-26 \%$ ) for family-size versus $-19 \%(95 \% \mathrm{CI}=-21,-16 \%)$ for individual-size beverages. <br> 3. There was no significant change in volume sold of untaxed beverages in Cook County and its border area. |  |
| 26 |  | 1. Volume sold of taxed beverage fell, on average, by $22 \%(p<0.001)$ in the first year following the implementation of a soda tax. <br> 2. Volume sold of taxed beverages fell to a greater extent for family-versus individual-size beverages ( $31 \%$ vs. $10 \%$ ) and fell to a greater extent for soda ( $29 \%$ ) compared to all other beverage types. <br> 3. Moderate substitution to untaxed beverages was found-volume sold of untaxed beverages increased by $4 \%$ ( $p<0.05$ ). <br> 4. There was no significant increase in the overall volume sold of taxed beverages in the 2 -mile border area of Seattle. | On average, in the first year of post-tax implementation, prices of taxed beverages rose by $1.03 ¢ / \mathrm{oz}$. ( $\mathrm{p}<0.001$ ), corresponding to a $59 \%$ tax passthrough rate. |
| 27 |  |  | 1. There was a $119 \%$ tax pass-through rate across all taxed beverages in Cook County compared to its comparison site. <br> 2. This price increase was $34 \%$ for taxed beverages. <br> 3. For untaxed beverages, prices increased slightly by $0.04 \AA /$ oz. driven mainly by an increase in milk prices ( $0.12 \phi / \mathrm{oz}$ ). <br> 4. Pass-through was higher for individual-size (126\%) compared to family-size (117\%) beverages and higher for energy drinks (145\%) compared to other sweetened beverages. <br> 5. Based on the baseline prices of different categories and sizes of beverages, the effective percentage increase in beverage prices ranged from a $52 \%$ increase for family-size soda to a $10 \%$ increase for family-size energy drinks. |
| 28 | People in Philadelphia were more likely to decrease their frequency of SSB consumption ( $39.2 \%$ vs. $33.5 \%$ ), and less likely to increase their frequency of SSB consumption ( $38.9 \%$ vs. $43.0 \%$ ) relative to those residing in untaxed cities. |  |  |

TABLE 3 (Continued)

| Study ID | Effects of a soda tax on consumption | Effects of a soda tax on sales or purchases | Effects of a soda tax on prices |
| :---: | :---: | :---: | :---: |
| 29 |  | 1. Purchases of taxed beverages declined by 6.1 floz . ( $95 \% \mathrm{CI}=-9.9,-2.4$; $\mathrm{p}<0.001$ ), corresponding to a $42 \%$ decline in Philadelphia compared with Baltimore; there were no significant changes in purchases of nontaxed beverages. <br> 2. Although there was no significant moderation by neighborhood income or customer education level, declines in taxed beverage purchases were larger among customers shopping in low-income neighborhoods ( -7.1 fl oz.; $95 \% \mathrm{CI}=-13.0,-1.1 ; \mathrm{p}=0.001$ ) and individuals with lower education levels ( -6.9 floz .; $95 \% \mathrm{CI}=-12.5,-1.3 ; p=0.001$ ). | Taxed beverage prices increased $2.06 \$ / \mathrm{oz}$. ( $95 \% \mathrm{CI}=1.75,2.38 ; \mathrm{p}<0.001$ ), with $137 \%$ of the tax passed through to prices 2 years after tax implementation in Philadelphia compared to Baltimore |
| 30 |  |  | The tax was largely, but not completely, passed through to consumers. In both the hand-collected store data and restaurant data, pass-through is slightly less than $75 \%$, whereas pass-through is just over $50 \%$ using the scanner data; consumers bear most, but not all, of the largest tax on sugar-sweetened beverages. |
| 31 | 1. Intake of soda reduced by 0.81 servings per week in Philadelphia compared with all other comparison cities 2 years after tax implementation. <br> 2. There was no significant difference in $100 \%$ juice or milk intake. <br> 3. In subgroup analyses, the tax was associated with a reduction of 1.13 servings per week in Hispanic/Latinx adolescents ( $95 \% \mathrm{CI}=-2.04,-0.23$ servings; $p=0.01$ ) and 1.2 servings per week in adolescents with obesity ( $95 \% \mathrm{CI}=-2.33,-0.13$ servings; $\mathrm{p}=0.03$ ). |  |  |
| 32 |  | Taxed beverages sales fell by $14 \%$, but $46 \%$ of this decrease is offset with an increase in the border area. | The taxed beverage prices increased by $0.49 \not \subset / \mathrm{oz}$. in Oakland relative to Sacramento. Tax pass-through is $49 \%$. |
| 33 |  |  | Taxed beverage prices increased by $0.73 \mathrm{\phi} / \mathrm{oz}$. ( $95 \% \mathrm{CI}=0.47,1.00$ ) on average in supermarkets and grocery stores in Oakland relative to Sacramento and $0.74 \not \subset / \mathrm{oz}$. ( $95 \% \mathrm{CI}=0.39,1.09$ ) in pharmacies, but did not change in convenience stores ( -0.09 $\phi / \mathrm{oz} ., 95 \% \mathrm{CI}=-0.56,0.39$ ). Untaxed beverage prices overall increased by $0.40 \phi / \mathrm{oz}$. ( $95 \% \mathrm{CI}=0.05,0.75$ ) in pharmacies but did not change in other store types. Prices of taxed individual-size soda specifically increased in all store types, by 0.91-2.39 $\phi / \mathrm{oz}$. ( $\mathrm{p}<0.05$ ), as did prices of untaxed individual-size soda in convenience stores ( 0.79 \&/ oz., $95 \% \mathrm{CI}=0.01,1.56$ ) and pharmacies ( 1.66 \$/oz., $95 \% \mathrm{CI}=0.09,3.23$ ). |
| 34 |  |  | 1. In fast-food restaurants, the price of bottled regular soda increased by $1.44 \not \subset /$ oz. $(95 \% \mathrm{CI}=0.50,2.73)$, with tax passthrough rate of $144 \%$. <br> 2. The price of bottled diet soda increased by 1.17 cents/oz. $(95 \% \mathrm{CI}=0.07,2.13)$. <br> 3. There were not statistically significant price effects for unsweetened beverages and fountain drinks. |


| Study ID | Effects of a soda tax on consumption | Effects of a soda tax on sales or purchases | Effects of a soda tax on prices |
| :---: | :---: | :---: | :---: |
| 35 |  | 1. Taxed beverages sales fell by $22 \%$ in Seattle relative to Portland. <br> 2. Declines were larger for familysize ( $29 \%$ ) compared to individual-size (10\%) beverages; particularly for soda ( $36 \%$ decrease for family-size compared to no change for individual-size). <br> 3. There was no change in volume sold of taxed beverages in Seattle's <br> 2-mile border area, suggesting no cross-border shopping. | Prices of taxed beverages increased by 1.04 ¢/oz. 2-year post-tax, with $59 \%$ tax pass-through rate. |
| 36 |  | A $5 \%$ volume reduction in Washington but fail to detect an effect in Berkeley. | Prices in Washington reacted sharply and promptly (often by a larger magnitude than the tax), whereas retail prices in Berkeley reacted marginally (by less than $30 \%$ the magnitude of the tax). |
| 37 |  | 1. The total volume of taxed beverages per store sold in Philadelphia decreased by $46 \%$. <br> 2. A large amount of cross-shopping to stores outside of Philadelphia off-sets more than half of the reduction in sales in the city and reduces the net decrease in sales of taxed beverages to only $22 \%$. | The tax is passed through at an average rate of $97 \%$, leading to a $34 \%$ price increase. |
| 38 |  | Each cent per ounce of taxes causes the sales quantity of taxed beverages to decrease in a range of 5.1-14.4\%. | Each cent per ounce of taxes causes the price of the taxed beverages to increase in a range from 0.47 to $0.98 \$ / \mathrm{oz}$. |
| 39 |  | 1. Taxed beverage volume sold decreased by $18 \%$ in Oakland relative to Sacramento, with a larger decrease for family-size beverages (23\%) relative to individual-size beverages ( $8 \%$ ). <br> 2. There was a $9 \%$ increase in volume sold of taxed beverages in the twomile border area surrounding Oakland relative to the Sacramento border area, driven by a $12 \%$ increase for family-size taxed beverages. <br> 3. After accounting for this cross-border shopping, there was a net decrease of 6\% in taxed beverage volume sold in Oakland. <br> 4. There was no significant change in untaxed beverage volume sold in either Oakland or its border area relative to their respective comparison sites. | Taxed beverage prices increased by $0.67 \$ / \mathrm{oz}$. in Oakland relative to Sacramento, corresponding to $67 \%$ pass-through. |
| 40 |  | 1. Taxed beverage volume sales in stores decreased by $50 \%$ ( $95 \% \mathrm{CI}=36$, $61 \%$ ), volume sales of nontaxed beverages did not change after tax implementation. <br> 2. After accounting for cross-border shopping, taxed beverage volume sales decreased in Philadelphia by $35 \%$ in 2018. Volume sales of nontaxed beverage concentrates increased on average by $34 \%$, but there was no evidence of substitution to high-calorie foods. | After tax implementation, taxed beverage prices in Philadelphia increased by 1.02 ¢/oz. ( $95 \% \mathrm{CI}=0.94,1.11 ; 68 \%$ pass through). |

(81). Leider et al. estimated that two-thirds of the overall decrease in the volume sold of taxed beverages in Oakland was offset by crossborder shopping (80). Léger and Powell estimated that $46 \%$ of Oakland's taxed beverage sales decrease was offset by cross-border shopping (77). Powell and Leider reported that there was no significant change in taxed beverage sales in Seattle's two-mile border area, suggesting there was no cross-border shopping for taxed beverages (79).

### 3.4. Meta-analysis

Table 4 summarizes the results from meta-analyses. Data from 27 studies contributed to the meta-analysis: 22 focused on the change in prices of taxed beverages and tax passthrough rate (37-$39,42,45,48-51,68-73,75-81), 11$ on purchases of taxed beverages ( $41,42,48,49,69,73,74,77,79-81$ ), three on the change in consumption of taxed beverages $(46,47,52), 15$ on the change
in prices of untaxed beverages $(38,39,42,45,49-51,68,71,73$, $75-78,81$ ), and five on the change in the purchases of untaxed beverages $(48,74,77,80,81)$. Soda tax implementation was found to be associated with an increase in the prices of taxed beverages by $1.06 \not \subset / \mathrm{oz} .\left(95 \% \mathrm{CI}=0.90,1.22 ; I^{2}=98.8 \%\right.$; RE), a reduction in the purchases of taxed beverages by $27.3 \% ~(95 \% \mathrm{CI}=19.3,35.4 \%$; $\left.I^{2}=97.2 \% ; \mathrm{RE}\right)$. The soda tax passthrough rate was estimated to be $79.7 \%$ ( $95 \% \mathrm{CI}=65.8,93.6 \% ; I^{2}=99.1 \% ; \mathrm{RE}$ ). Following the soda tax implementation, the prices of untaxed beverages were estimated to increase by $0.08 ¢ /$ oz. ( $95 \% \mathrm{CI}=0.04,0.12 ; I^{2}=77.3 \%$; RE). No changes were observed for the purchases of untaxed beverages and the consumption of taxed beverages $(p>0.05)$. Meta-regression found that $1 \phi /$ oz. increase in the soda tax rate was associated with an increase in the prices of taxed beverages by $0.84 ¢ / \mathrm{oz}$. $(95 \%$ $\mathrm{CI}=0.33,1.35$ ). No dose-response effect of soda taxes was found for purchases and consumption of taxed beverages or the tax passthrough rate ( $p s>0.05$ ). No publication bias was identified by Egger's or Begg's tests ( $p s>0.05$ ).

TABLE 4 Results from meta-analyses and publication bias tests.

| Outcome | Studies included in the metaanalysis | $I^{2}$ index | Pooled Effect size (95\% CI) | Model | Publication bias test |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | $p$-value Egger's test | $p$-value Begg's test |
| Price change in taxed beverages ( $\phi / \mathrm{oz}$ ) | Falbe et al. (45), Cawley and Frisvold (37), Silver et al. (51), Cawley et al. $(38,39)$, Roberto et al. (49), Bleich et al. (69), Cawley et al. (42), Falbe et al. (70), JonesSmith et al. (71), Marinello et al. (72), Powell and Leider (73), Powell and Leider (48), Powell et al. (75), Bleich et al. (76), Léger and Powell (77), Marinello et al. (78), Powell and Leider (79), Seiler et al. (50), Leider and Powell (80), Petimar et al. (81) | 98.8\% | 1.06 (0.90, 1.22) | RE | 0.75 | 0.76 |
| Price change in untaxed beverages ( $\phi / \mathrm{oz}$ ) | Falbe et al. (45), Silver et al. (51), Cawley et al. $(38,39)$, Roberto et al. (49), Seiler et al. (50), Cawley et al. (42), Jones-Smith et al. (71), Powell and Leider (73), Powell et al. (75), Bleich et al. (76), Léger and Powell (77), Marinello et al. (78), Petimar et al. (81) | 77.3\% | 0.08 (0.04, 0.12) | RE | 0.08 | 0.55 |
| Purchase change in taxed beverages | Cawley et al. (41), Roberto et al. (49), Cawley et al. (42), Powell and Leider (73), Powell and Leider (48), Powell et al. (74), Bleich et al. (69), Léger and Powell (77), Powell et al. (53, 54), Leider and Powell (80), Petimar et al. (81) | 97.2\% | $\begin{gathered} -27.3 \%(-35.4 \\ -19.3 \%) \end{gathered}$ | RE | 0.98 | 0.39 |
| Purchase change in untaxed beverages | Powell et al. (74), Powell and Leider (48), Léger and Powell (77), Leider and Powell (80), Petimar et al. (81) | 0.0\% | 3.1\% (0.3, 5.9\%) | FE | 0.09 | 0.09 |
| Consumption change in taxed beverages | Falbe et al. (46), Zhong et al. (52), Lee et al. (47) | 48.7\% | $\begin{gathered} -36.1 \%(-51.4, \\ -20.8 \%) \end{gathered}$ | FE | 0.59 | 1.00 |
| SSB tax pass-through rate | Falbe et al. (45), Cawley and Frisvold (37), Silver et al. (51), Cawley et al. $(38,39)$, Roberto et al. (49), Bleich et al. (69), Cawley et al. (42), Falbe et al. (70), JonesSmith et al. (71), Marinello et al. (72), Powell and Leider (73), Powell and Leider (48), Powell et al. (75), Bleich et al. (76), Léger and Powell (77), Marinello et al. (78), Powell et al. $(53,54)$, Seiler et al. (50), Leider and Powell (80), Petimar et al. (81) | 99.1\% | $\begin{gathered} 79.7 \% ~(65.8, \\ 93.6 \%) \end{gathered}$ | RE | 0.44 | 0.22 |

[^0]
### 3.5. Study quality assessment

Table 5 reports criterion-specific and overall ratings from the study quality assessment. The included studies, on average, scored six out of ten, with scores ranging from 3 to 9 . All studies collected and analyzed both pre- and post-tax outcomes for the intervention group (i.e., retailers or consumers residing in the taxing jurisdiction) and included a control group (i.e., retailers or consumers living outside the taxing jurisdiction or beverages not subject to the soda tax). The majority of studies ( $n=36$ ) had comparable intervention and control groups. Additionally, most studies had a sample representative of the stores or the population residing in the taxing jurisdiction $(n=17)(35,36,38-42,44,45,50,51$, $60,70-72,74,75)$. Furthermore, most studies included the change in beverage prices $(n=28)(36-39,42,43,45,48-51,59,61-63,68-73$, $75-81)$, sales or purchases $(n=22)(35,36,40-42,44,48-51,56,57,62$, $63,69,73,74,76,77,79-81)$ and consumption $(n=11)$ as outcomes (36, $40,42,44,46,47,50-52,58,60)$. Additionally, several studies assessed the impact of cross-border sales or purchases $(n=17)(36,37,39,40,42$, $46,48-50,57,62,69,74,77,79-81)$. However, we noted some limitations in the quality of the included studies. Only two studies provided a sample size justification ( 58,62 ), and nine studies performed adequate statistical procedures to adjust for temporal trends of the outcomes ( $37,44,47,49$, $50,52,60,73,79)$. To enhance the rigor of future research in this area, we recommend that studies provide clear justifications for sample sizes and utilize appropriate statistical procedures to account for temporal trends. Additionally, researchers should aim to improve reporting practices, including the transparent reporting of study limitations and potential sources of bias.

## 4. Discussion

A few previous meta-analyses estimated the effect of soda taxes or SSB prices on beverage sales, purchases, and consumption. Nakhimovsky et al. reviewed studies in middle-income countries and estimated that a $10 \%$ increase in SSB prices was associated with a reduction in SSB consumption from 1.2-9.3 kcal per person per day (83). Teng et al. estimated that a $10 \%$ soda tax was associated with a decline in SSB purchases and consumption by $10 \%$ (84). Escobar et al. (14) and Powell et al. (85) estimated the price-elasticity of SSBs to be -1.30 and -1.21 , indicating that a $10 \%$ increase in SSB prices was associated with a reduction in demand for SSBs by 13.0 and $12.1 \%$, respectively. Afshin et al. (12) reported that a $10 \%$ increase in SSB prices was associated with a reduction in SSB consumption by $7 \%$. Those estimates from the previous meta-analyses were comparable to the estimated impact of soda taxes on SSB purchases and consumption in this review. However, the key differences are: previous reviews and meta-analyses were nearly exclusively based on the "proxy" or "modeling" studies due to lacking a soda tax in reality and collected data from multiple other countries besides the US. This review contributed to the literature by comprehensively assessing the efficacy of soda taxes based on evidence reported exclusively from natural experiments (the "local" studies). In addition, the review provided quantitative estimates of the magnitude of the tax effect. Review findings can potentially inform policymakers in designing and implementing soda taxes to curb the obesity epidemic.

Several main criticisms of soda taxes are worth noting. First, substitution effects in addition to cross-border shopping may
partially offset the intended effect of soda taxes on SSB purchases and consumption ( 49,50 ). A nationwide soda tax adoption could eliminate cross-border shopping (49). Second, the soda tax could be regressive and the impact of soda consumption is greatest in low income populations. However, a systematic review revealed only modest differences- $0.1-1.0 \%$ and $0.03-0.60 \%$ of annual household income paid for soda taxes for low- and high-income households, respectively (86). Third, revenues collected by soda taxes were rarely earmarked for healthy diet promotion, such as in the form of healthy food subsidies paid to low-income households (87). Fourth, relatively arbitrary classifications of SSBs result in taxing some low-sugar beverages but exempting some high-sugar ones (88). The beverage industry has used these policy loopholes to lobby against soda tax bills (88). Finally, some researchers questioned the efficiency of the volumetric soda tax and suggested taxing the amount of sugar in a drink rather than the volume of liquid that accompanied the sugar to boost the soda tax's health benefits and overall economic gains (89).

Despite the various criticisms of soda taxes, it remains a potentially effective policy leverage to nudge people toward reducing purchases and consumption of SSBs. There is also preliminary evidence linking soda tax to reduced body weight, BMI, and overweight or obesity (14, 15). The revenues collected from soda taxes could support other health interventions, such as nutrition education campaigns or healthy food subsidies, which may reinforce the long-term sustainability of soda taxes on dietary pattern changes. Such policy arrangements might also alleviate health disparities at the population level.

This study comprehensively reviews evidence reported from natural experiments on soda taxes implemented across US local taxing jurisdictions. The strengths of the included studies encompassed realworld policy interventions providing causal inferences, the inclusion of comparison cities, an elucidation of the local soda tax implementation processes, and longitudinal changes in price, consumption, and purchases pre-and post-tax. However, a few limitations of this review and the included studies should be noted. First, all studies were natural experiments focused on one or a few US cities implementing a soda tax. The limited geographic coverage and nationally non-representative sample have confined the generalizability of study findings. Therefore, caution should be exercised when extrapolating the results to other cities or jurisdictions. Second, DID is a quantitative method to reveal causal references, but the resulting evidence should not be overstated due to lacking a randomization design (82). Third, it is important to consider the magnitude of the tax rates examined in the included studies. The range of tax rates observed in the reviewed studies was relatively narrow, typically between 1 and $2 \phi / \mathrm{oz}$. Thus, the effects reported in the included studies may not necessarily apply to different tax magnitudes. Non-linearities in the effects on tax pass-through or purchases may exist, and these non-linearities could be influenced by factors such as company profit considerations or consumer price sensitivity. Additionally, we acknowledge that the slight increase in prices of untaxed beverages following soda tax implementations is an interesting finding. However, the reasons behind this increase are not entirely clear from available evidence. It could be a result of temporal changes in the prices of untaxed beverages or it may reflect substitution effects, where higher demand for untaxed beverages leads to a price increase. Further research is needed to understand the underlying factors contributing to this observation.

TABLE 5 Study quality assessment.

| Criterion Study ID | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 | 16 | 17 | 18 | 19 | 20 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1. Did the study collect and analyze both pre- and posttax outcomes for the intervention group (i.e., retailers or consumers residing in the taxing jurisdiction)? | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 2. Did the study include a control group (i.e., retailers or consumers residing outside of the taxing jurisdiction, or beverages not subject to soda taxes)? | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 3. Were the intervention and control groups similar in all aspects except for their soda tax implementation status? If not, were adequate statistical procedures performed to adjust for group differences? | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 0 | 1 | 1 | 1 | 1 |
| 4. Were adequate statistical procedures performed to adjust for the temporal trends of the outcomes? | 0 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 0 |
| 5. Was a sample size justification provided? | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6. Was the study sample representative of the stores or the population residing in the taxing jurisdiction? | 1 | 0 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 1 | 1 |
| 7. Was the change in beverage prices following the soda tax implementation included as an outcome? | 1 | 0 | 1 | 0 | 1 | 1 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 1 | 1 | 1 | 1 |
| 8. Was the change in beverage sales or purchases following the soda tax implementation included an outcome? | 0 | 0 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 1 | 1 | 0 | 1 | 1 | 1 | 1 | 0 | 0 |
| 9. Was the change in beverage consumption following the soda tax implementation included as an outcome? | 0 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 1 | 0 | 0 | 1 | 1 | 0 | 0 | 0 | 1 | 0 | 0 |
| 10. Did the study assess the impact of cross-border sales or purchases following the soda tax implementation? | 0 | 1 | 1 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 1 | 0 | 1 | 1 | 1 | 0 |
| Total score | 5 | 5 | 6 | 7 | 7 | 8 | 4 | 4 | 4 | 5 | 5 | 5 | 7 | 5 | 7 | 3 | 6 | 8 | 6 | 5 |

TABLE 5 (Continued)

| Criterion Study ID | 21 | 22 | 23 | 24 | 25 | 26 | 27 | 28 | 29 | 30 | 31 | 32 | 33 | 34 | 35 | 36 | 37 | 38 | 39 | 40 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1. Did the study collect and analyze both pre- and post-tax outcomes for the intervention group (i.e., retailers or consumers residing in the taxing jurisdiction)? | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 2. Did the study include a control group (i.e., retailers or consumers residing outside of the taxing jurisdiction, or beverages not subject to soda taxes)? | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 3. Were the intervention and control groups similar in all aspects except for their soda tax implementation status? If not, were adequate statistical procedures performed to adjust for group differences? | 1 | 1 | 1 | 1 | 0 | 1 | 1 | 0 | 1 | 1 | 0 | 1 | 1 | 0 | 1 | 1 | 1 | 1 | 1 | 1 |
| 4. Were adequate statistical procedures performed to adjust for the temporal trends of the outcomes? | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 1 | 0 | 1 | 0 | 0 | 0 |
| 5. Was a sample size justification provided? | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 6. Was the study sample representative of the stores or the population residing in the taxing jurisdiction? | 1 | 0 | 1 | 0 | 1 | 0 | 1 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| 7. Was the change in beverage prices following the soda tax implementation included as an outcome? | 1 | 0 | 1 | 1 | 0 | 1 | 1 | 0 | 1 | 1 | 0 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 |
| 8. Was the change in beverage sales or purchases following the soda tax implementation included an outcome? | 0 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 1 | 0 | 0 | 1 | 0 | 0 | 1 | 1 | 1 | 1 | 1 | 1 |
| 9. Was the change in beverage consumption following the soda tax implementation included as an outcome? | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| 10. Did the study assess the impact of cross-border sales or purchases following the soda tax implementation? | 0 | 1 | 0 | 0 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 1 | 1 | 1 | 0 | 1 | 1 |
| Total score | 5 | 5 | 5 | 6 | 5 | 6 | 5 | 4 | 5 | 4 | 5 | 6 | 4 | 3 | 7 | 7 | 9 | 5 | 6 | 6 |

For each criterion, a score of one was assigned if "yes" was the response, whereas a score of zero was assigned otherwise. A study-specific global score ranging from 0 to 10 was calculated by summing up scores across all criteria. The study quality assessment helped measure the strength of scientific evidence but was not used to determine the inclusion of studies.

## 5. Conclusion

This study systematically reviewed evidence from natural experiments regarding the impact of soda taxes on beverage prices,
sales, purchases, and consumption in the US. Soda tax implementation was associated with increased prices of taxed beverages and reduced purchases and consumption of taxed beverages. Soda taxes could be effective policy leverage to nudge people toward purchasing and
consuming fewer SSBs. Future research should examine evidencebased classifications of SSBs, more targeted use of revenues generated by taxes to reduce health and income disparities, and the feasibility of redesigning the soda tax to improve its efficiency.

## Data availability statement

The original contributions presented in the study are included in the article/Supplementary material, further inquiries can be directed to the corresponding author.

## Author contributions

JS, JW, and RA designed the study, wrote the manuscript, and revised the manuscript. JW, FY, and JS jointly designed the search algorithm and screened articles. JS and JW performed data extraction and constructed the summary tables. All authors contributed to the article and approved the submitted version.

## Funding

This research was funded by the Fundamental Research Funds for the Central Universities, CUGB, grant number 2-9-2020-036; The

Ministry of Education of Humanities and Social Science project, grant number 19YJC890044.

## Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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## Supplementary material

The Supplementary material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fpubh.2023.1126569/ full\#supplementary-material

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[^0]:    RE: random-effect model, FE: fixed-effect model.

