University of Groningen

Effects of race and ethnicity on perinatal outcomes in high-income and upper-middle-income countries

IPPIC Collaborative Network; Sheikh, Jameela; Allotey, John; Kew, Tania; Fernández-Félix, Borja M.; Zamora, Javier; Khalil, Asma; Thangaratinam, Shakila

IMPORTANT NOTE: You are advised to consult the publisher's version (publisher's PDF) if you wish to cite from it. Please check the document version below.

Document Version
Publisher's PDF, also known as Version of record

Publication date:
2022

Link to publication in University of Groningen/UMCG research database

Effects of race and ethnicity on perinatal outcomes in high-income and upper-middle-income countries: an individual participant data meta-analysis of 2 198 655 pregnancies

Jameela Sheikh*, John Allotey*, Tania Kew, Borja M Fernández-Félix, Javier Zamora, Asma Khalil, Shakila Thangaratinam, IPPIC Collaborative Network†

Summary

Background Existing evidence on the effects of race and ethnicity on pregnancy outcomes is restricted to individual studies done within specific countries and health systems. We aimed to assess the impact of race and ethnicity on perinatal outcomes in high-income and upper-middle-income countries, and to ascertain whether the magnitude of disparities, if any, varied across geographical regions.

Methods For this individual participant data (IPD) meta-analysis we used data from the International Prediction of Pregnancy Complications (IPPIC) Network of studies on pregnancy complications; the full dataset comprised 94 studies, 53 countries, and 4 539 640 pregnancies. We included studies that reported perinatal outcomes (neonatal death, stillbirth, preterm birth, and small-for-gestational-age babies) in at least two racial or ethnic groups (White, Black, south Asian, Hispanic, or other). For our two-step random-effects IPD meta-analysis, we did multiple imputations for confounder variables (maternal age, BMI, parity, and level of maternal education) selected with a directed acyclic graph. The primary outcomes were neonatal mortality and stillbirth. Secondary outcomes were preterm birth and a small-for-gestational-age baby. We estimated the association of race and ethnicity with perinatal outcomes using a multivariate logistic regression model and reported this association with odds ratios (ORs) and 95% CIs. We also did a subgroup analysis of studies by geographical region.

Findings 51 studies from 20 high-income and upper-middle-income countries, comprising 2 198 655 pregnancies, were eligible for inclusion in this IPD meta-analysis. Neonatal death was twice as likely in babies born to Black women than in babies born to White women (OR 2·00, 95% CI 1·44–2·78), as was stillbirth (2·16, 1·46–3·19), and babies born to Black women were at increased risk of preterm birth (1·65, 1·46–1·88) and being small for gestational age (1·39, 1·13–1·72). Babies of women categorised as Hispanic had a three-times increased risk of neonatal death (OR 3·34, 95% CI 2·77–4·02) than did those born to White women, and those born to south Asian women were at increased risk of preterm birth (OR 1·26, 95% CI 1·07–1·48) and being small for gestational age (1·61, 1·32–1·95). The effects of race and ethnicity on preterm birth and small-for-gestational-age babies did not vary across regions.

Interpretation Globally, among underserved groups, babies born to Black women had consistently poorer perinatal outcomes than White women after adjusting for maternal characteristics, although the risks varied for other groups. The effects of race and ethnicity on adverse perinatal outcomes did not vary by region.

Funding National Institute for Health and Care Research, Wellbeing of Women.

Introduction

Inequalities and inequities faced by pregnant women contribute to the differential rates of adverse perinatal outcomes.[1–3] In countries with multiethnic populations, the variations in perinatal outcomes between different racial and ethnic groups reflect the underlying health inequalities in maternity care.[4] This difference has an impact on the health of future generations in the short and long term.[5]

To date, studies documenting racial and ethnic inequality and poor maternal and offspring outcomes have focused on specific groups of women,[6] or had a country-specific focus.[7,8] This makes it challenging to investigate the degree of inequality and inequities faced

globally by women from various underserved and under-represented racial and ethnic groups. The disparities in pregnancy outcomes are particularly stark in high-income and upper-middle-income countries, where the overall quality of health care is high and mortality rates are low. In the UK, the *Mothers and Babies: Reducing Risk through Audits and Confidence Enquiries across the UK* report on confidential enquiries showed that the rates of neonatal death and stillbirth in babies of Black and Asian women are double those of White women.[1,9] Similar trends are seen in the USA, with high rates of preterm birth and low birthweight in babies of Black women compared to babies of White women.[10]

Research in context

Evidence before this study

Individual studies done in countries with multiethnic populations suggest associations between underserved racial and ethnic backgrounds and adverse perinatal outcomes. Studies reporting the effects of race and ethnicity on adverse perinatal outcomes generally do not isolate the causal effect of race and ethnicity by adjusting for other factors such as socioeconomic background and health conditions. We did a MEDLINE search with no language restrictions from database inception to Jan 31, 2022, for systematic reviews on race and ethnicity and adverse perinatal outcomes, using the search terms "ethnicity" OR "race" AND "neonatal mortality" OR "stillbirth" OR "preterm" OR "SGA" OR "small for gestation". Two systematic reviews analysing the relationship between race and ethnicity and preterm birth reported increased risks of preterm birth in Black women (odds ratio 1·5–2·0) compared to non-Black and White women, mostly from studies done in the USA. To the best of our knowledge, no study has so far provided a global overview of the effect of race and ethnicity on neonatal deaths, stillbirth, preterm birth, and small-for-gestational-age babies, and whether this effect varies by region.

Added value of this study

In this meta-analysis we provided a global outlook of the magnitude of the association between race and ethnicity and adverse perinatal outcomes across high-income and upper-middle-income countries. Our individual participant data meta-analysis of more than two million pregnancies from multiple cohorts worldwide showed that Black women are at higher risk of adverse perinatal outcomes of neonatal death, stillbirth, preterm birth, and small-for-gestational-age babies than White women, even after adjusting for maternal characteristics. These racial disparities in perinatal outcomes are consistently observed across all geographical regions. Our study is, to the best of our knowledge, the first to assess the effect of race and ethnicity on perinatal outcomes across high-income and upper-middle-income countries.

Implications of all the available evidence

The disparities and inequalities in pregnancy outcomes observed in women from underserved and under-represented racial and ethnic groups across geographical regions highlight the need for a global approach to this problem. We require a holistic approach that complements multifaceted antenatal interventions, with a life course approach tackling race-related and ethnicity-related barriers faced by girls and young women, particularly Black women, who are the most affected. Race and ethnicity data, and relevant confounders (eg, maternal education), should be routinely collected in detail alongside qualitative evaluations, to identify the magnitude of the risks faced by women in various racial and ethnic subgroups and plan appropriate interventions for those with the highest need.

Research into the causal relationship of race and ethnicity with adverse health outcomes is challenging, particularly when considering regression models investigating the effects of race and ethnicity. Consideration of causal pathways and the relationship between variables is crucial to isolate the causal effect of race and ethnicity, a social construct present before the index pregnancy, on perinatal outcomes by controlling for other confounding variables.[11] This approach of using a causal pathway when investigating perinatal outcomes challenges the certainty of the degree of influence a woman's underlying socioeconomic background and health status, and the health-care system, has on her clinical outcome, compared to the inequalities related to race and ethnicity.

We aimed to quantify the effects of race and ethnicity in women from underserved groups in high-income and upper-middle-income countries on neonatal deaths and stillbirths primarily, and on preterm births and small-for-gestational-age babies secondarily, after adjusting for confounders in the causal pathway. We also aimed to determine the variations in the effects of race and ethnicity on offspring outcomes across studies from various geographical regions.

Methods

Search strategy and selection criteria

Our individual participant data (IPD) meta-analysis was based on a prospectively registered protocol,[12] and we reported our findings in accordance with the Preferred Reporting Items for Systematic Review and Meta-Analyses of Individual Participant Data (PRISMA-IPD) statement (appendix pp 1–3).[13]

Eligible studies were identified from the International Prediction of Pregnancy Complications (IPPIC) IPD Network without any language restrictions (figure 1).[14,15] The studies in the network were identified by searching major databases such as MEDLINE, Embase, Cochrane (Wiley) CENTRAL, Science Citation Index (Web of Science), CINAHL (EBSCO), the ISRCTN Registry, UK Clinical Trials Gateway, WHO International Clinical Trials Portal, and ClinicalTrials.gov, specialist abstract and conference proceeding resources (British Library's ZETOC and Web of Science Conference Proceedings Citation Index) for outcomes such as pre-eclampsia, fetal growth restriction, and birthweight (from database inception to August, 2019). Details of the search, identification, inclusion of studies, and IPD harmonisation for the IPPIC database are provided elsewhere.[15,16] The IPPIC dataset (comprising 53 countries, 94 studies, and 4 539 640 pregnancies) contains IPD from observational studies and cohorts nested within randomised studies reporting various maternal and perinatal outcomes.[16] Studies obtained data on race and ethnicity through various methods, including self-reporting by the woman, routine data collected in medical records, or as recorded by the research team with prespecified definitions.

We included IPD on singleton pregnancies in the IPPIC dataset providing data on adverse perinatal outcomes (neonatal death, stillbirth, preterm birth, and small-for-gestational-age babies) in at least two racial and ethnic groups. We included only studies from high-income and upper-middle-income countries, as per the World Bank classification.[17] Although race and ethnicity can be defined separately, they are overlapping concepts and are often used interchangeably.[18] Therefore, throughout the Article, we use the terms race and ethnicity in line with current recommendations,[19] acknowledging that these are social constructs. In the IPPIC dataset, we harmonised the various definitions used to define the race and ethnicity of participants as White, Black, south Asian, Hispanic, and other groups (including those of multiracial, multiethnic, and east Asian origin). Black women comprised those of African origin, including African American and African Caribbean women; the south Asian group comprised women from the Indian subcontinent; and the Hispanic group comprised women in the USA of Spanish-speaking or Latin American descent or heritage. We categorised all women of Hispanic identity, irrespective of their racial identity (White Hispanic, Black Hispanic, Asian Hispanic, and other Hispanic) as a single group.[19,20] We considered Black, south Asian, Hispanic, and other populations to be underserved as reflected in the disparities and inequalities in health outcomes.[19]

We extracted data on women's characteristics such as age in years, parity (ie, nulliparous or multiparous), the highest level of maternal education attained (primary, secondary, or tertiary education), BMI, pre-existing or new-onset diabetes or hypertension, renal disease, autoimmune disease, and previous obstetric history of stillbirth or preterm birth. The primary outcomes were neonatal mortality (first 28 days of life)[21] and stillbirth (≥20 weeks' gestation).[22] Secondary outcomes were preterm birth (<37 weeks' gestation)[23] and small-for-gestational-age baby (birthweight <tenth centile).[24]

Two independent reviewers (JS and TK) assessed the methodological quality of the included studies by use of the Newcastle Ottawa Scale for selection, comparability, and outcome ascertainment bias.[25] Studies were considered to have a low risk of bias if they achieved four stars for selection, two for comparability, and three for ascertainment of the outcome. Studies achieving two or three stars for selection, one for comparability, and two for outcome ascertainment, were considered to have a medium risk of bias. When studies achieved one star for selection or outcome ascertainment, or zero stars for any of the three categories, this was regarded as a high risk of bias. The summary risk of bias for the study was determined by the total number of stars, where seven to nine stars was considered low risk, four to six was considered medium risk, and less than four stars meant the study was regarded as having a high risk of bias.
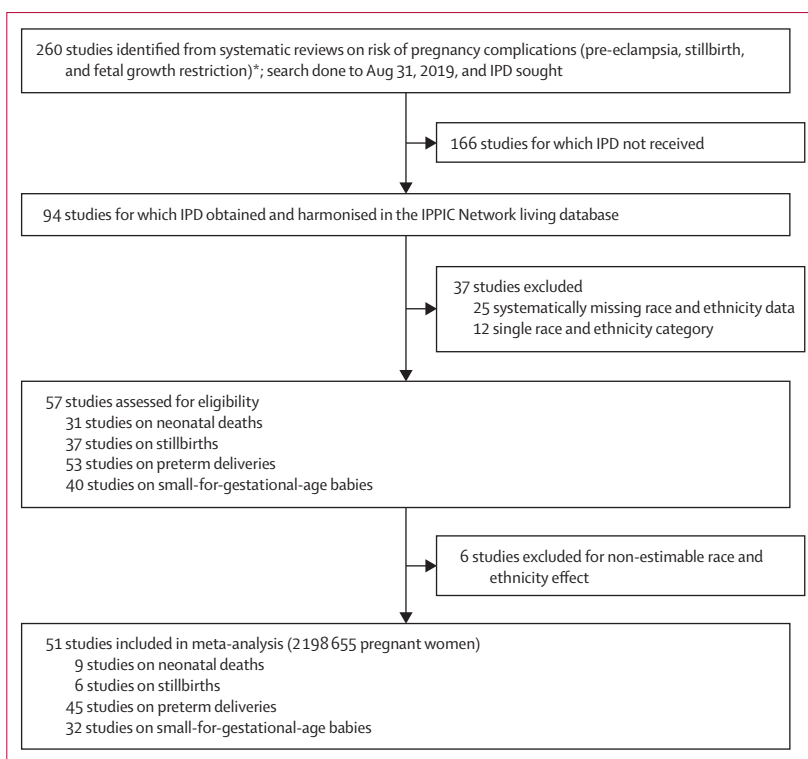
Figure 1: Study selection
IPD=individual participant data. IPPIC=International Prediction of Pregnancy Complications. *Townsend et al,[14] Allotey et al,[15] and Allotey et al (unpublished).

Data analysis

We did a two-step random-effects IPD meta-analysis. First, we did multiple imputations assuming a missing-at-random mechanism, and chained equations were used to generate 100 imputed datafiles for each cohort. Linear regression models were used for imputing continuous variables, logistic regression for binary variables, and multinomial logistic regression (or predictive mean matching in the presence of convergence issues) for categorical variables. We used outcome data in the imputation models to impute missing data on confounders.[26] However, when estimating the effects of race and ethnicity, we did not consider the imputed outcomes.[27] We did not impute when the values were completely missing or when more than 50% were missing.

We estimated the effects of race and ethnicity on perinatal outcomes by comparing pregnant women from underserved groups with White women in each individual cohort by fitting a multivariate logistic regression model in each imputed dataset. We assessed collinearity by estimating the variance inflation factor for all models.[28] Our proposed causal diagram and the assumptions are shown in the directed acyclic graph (figure 2).[29] We considered the exposure to be race and ethnicity, and that it affects the main outcomes of stillbirth and neonatal death through the causal pathway either directly, or is mediated through gestational age at
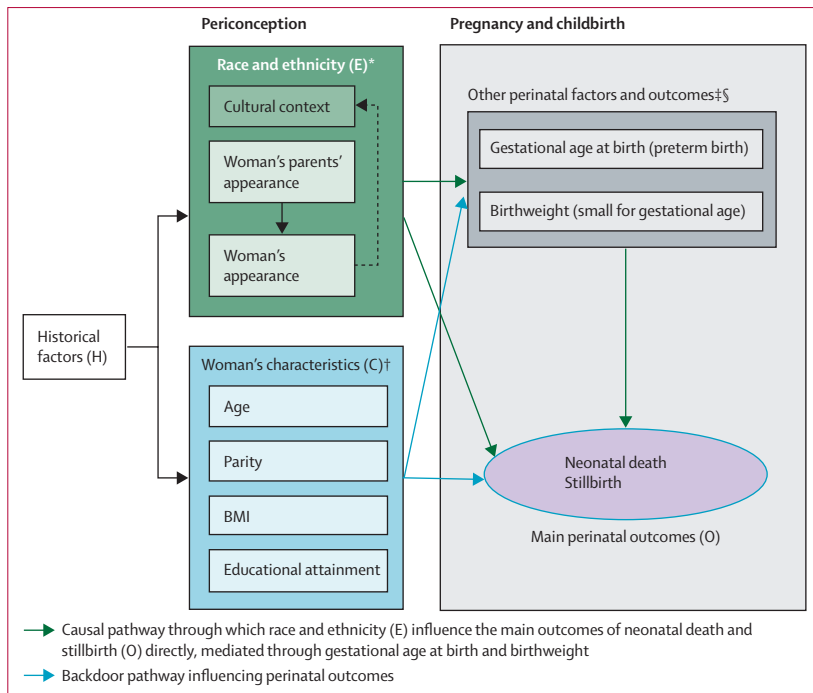
Figure 2: Causal diagram of race and ethnicity, and perinatal outcomes
The effects of race and ethnicity are considered to be the combined effects of the woman's appearance (phenotype including skin colour), her parents' appearance, and cultural context. The dashed line represents the potential influence of cultural context by the woman's appearance. Historical factors (H) include socioeconomic status of the family and neighbourhood at the time of the woman's conception. *Exposure. †Confounding is through a woman's characteristics such as age, parity, BMI, and educational attainment present at the time of her conception, which share a common history (H) with race and ethnicity (E); maternal educational attainment is a proxy for socioeconomic status. ‡Mediator is an intermediate variable between exposure (E) and outcome (O). §Collider is causally influenced by two or more variables.

used as the reference group. This process was repeated to obtain unadjusted (crude) estimations of the effects of race and ethnicity to check the impact of adjustments on the overall estimation of race and ethnicity effect.

We did subgroup analyses where appropriate by geographical region on the effects of race and ethnicity on perinatal outcomes. We classified regions as the USA and Canada; the UK; northern, western, and southern Europe (including France, Germany, Greece, Italy, Netherlands, Norway, and Spain) based on the UN geoscheme,[33] and other regions (including Australia, Brazil, and multi-country studies). We evaluated the robustness of our assumptions about missingness through several sensitivity analyses where we imputed the main exposure variable (race and ethnicity) under extreme scenarios of all missing cases being White women, then Black women, and so on. Sensitivity analyses were done by limiting the analysis to high-risk women, defined as those with risk factors such as previous stillbirth, previous preterm birth, pre-existing or new-onset diabetes or hypertension, maternal age older than 40 years, or obesity (BMI ≥30 kg/m²). Sensitivity analyses were also done to assess the impact of the imputation strategies by restricting the analysis to complete cases, and the impact of the study period on the observed effects of race and ethnicity through meta regression by year of recruitment (midpoint of recruitment period). Further sensitivity analysis were done by excluding one study that recruited women between 1959 and 1965;[34] and by excluding one multi-country study that involved women from low-income and middle-income countries.[35] All analyses were done with Stata (version 17).

Role of the funding source
The funders of the study had no role in study design, data collection, data analysis, data interpretation, writing of the manuscript, or the decision to submit the manuscript for publication.

birth and birthweight, which are proxies for the perinatal outcomes of preterm birth and a small-for-gestational-age baby. We also assumed that the woman's characteristics such as age, BMI, parity, and highest educational attainment that are present at the time of conception of her baby are related to complex historical factors (eg, family and neighbourhood socioeconomic status) present at the time of her own conception and birth.[11] As historical factors also have an influence on race-related and ethnicity-related factors, we used the woman's characteristics as confounding into the analysis to block the backdoor pathway from historical factors to perinatal outcomes. Since a woman's characteristics can also influence birthweight and gestational age at delivery (colliders), we refrained from adjusting for birthweight and gestational age in the analyses.[30] The effect of race and ethnicity was averaged over the imputed datasets by use of Rubin's rules within each cohort.[31]

In the second step, we used a random-effects model to pool the averaged effects estimated in the cohorts using the method of DerSimonian and Laird,[32] with the estimate of heterogeneity being taken from the Mantel-Haenszel model. Odds ratios (ORs) with 95% CIs were selected as the effect measure, with White women

Results
94 studies in the IPPIC IPD Network's database reported race and ethnicity, and adverse pregnancy outcomes. Of these, 51 provided the relevant IPD for 2 198 655 pregnant women from 20 high-income and upper-middle-income countries (figure 1; appendix p 4).

Of the 51 studies,[34–84] 42 were observational studies (35 prospective[34–45,47,48,50–54,56–61,63–66,70–75] and seven retrospective cohorts[46,49,55,62,67–69]), including six birth registries[48,63,73] and birth cohorts,[41,60,61] and nine were cohorts nested within randomised controlled trials[77–85] (appendix pp 5–20). Most studies were from the UK (13 studies),[43,53,59–61,64,73,74,77,78,80,81,83] followed by the USA (nine studies),[34,39,44,48,49,50,67,76,79] the Netherlands (six studies),[41,54,55,69,70,75] and Canada (four studies).[38,58,66,68] Regions were represented as follows: 17 studies were from northern, western and southern Europe; 13 were from the USA and Canada; 13 were from the UK; and eight were from other regions. 11 (22%) of the

	Maternal age	BMI	Parity	Educational attainment	Comparison		White women		Odds ratio (95% CI)	Weight (%)
					Events (n)	Women (N)	Events (n)	Women (N)		
Black women										
Stirrup et al (2015)[73]	√		√		22	7824	39	33283	2·39 (1·40–4·09)	12·66
Baschat et al (2014)[39]	√		√		31	803	13	775	2·57 (1·27–5·19)	10·21
Widmer et al (2015)[35]	√		√		36	761	109	2241	0·97 (0·66–1·44)	15·05
Jaddoe et al (2012)[41]	√	√	√	√	6	2146	10	4933	1·45 (0·46–4·57)	5·73
Caritis et al (1998)[76]	√	√	√	√	42	1772	16	1077	2·34 (1·26–4·37)	11·33
Mbah et al (2012)[48]	√	√		√	2881	247077	8093	1349397	1·82 (1·74–1·91)	19·05
Zhang et al (2001)[34]	√	√			9	502	13	1061	1·63 (0·66–4·02)	7·85
NICHD (2018)[49]	√	√	√		313	52518	240	116000	3·41 (2·85–4·06)	18·11
Overall					3340	313403	8533	1508767	2·00 (1·44–2·78)	100·00
I²=88·1%, p=0·000										
South Asian women										
Stirrup et al (2015)[73]	√		√		19	10393	39	33283	1·54 (0·88–2·68)	18·83
Widmer et al (2015)[35]	√		√		66	1450	109	2241	0·94 (0·68–1·30)	29·22
Huang et al (2010)[68]	√		√		6	35293	16	89350	0·96 (0·38–2·45)	9·32
Mbah et al (2012)[48]	√	√		√	8	1050	8093	1349397	1·47 (0·73–2·95)	14·34
NICHD (2018)[49]	√	√	√		39	9487	240	116000	1·83 (1·29–2·58)	28·29
Overall					138	57673	8497	1590271	1·33 (0·96–1·84)	100·00
I²=51·8%, p=0·081										
Hispanic women										
Caritis et al (1998)[76]	√	√	√	√	12	388	16	1077	2·48 (1·03–5·97)	4·46
NICHD (2018)[49]	√	√	√		210	40409	240	116000	3·38 (2·80–4·09)	95·54
Overall					222	40797	256	117077	3·34 (2·77–4·02)	100·00
I²=0·0%, p=0·499										
Other										
Widmer et al (2015)[35]	√		√		67	2857	109	2241	0·49 (0·36–0·67)	33·25
Mbah et al (2012)[48]	√	√		√	110	19466	8093	1349397	0·95 (0·79–1·16)	33·88
NICHD (2018)[49]	√	√	√		34	5512	240	116000	4·21 (2·92–6·07)	32·87
Overall					211	27835	8442	1467638	1·24 (0·44–3·49)	100·00
I²=97·5%, p=0·000										

√ Adjusted for that confounder

0·17   1·00   6·00

Decreased odds of neonatal death   Increased odds of neonatal death

Figure 3: Effect of race and ethnicity on the risk of neonatal deaths
Weights are from random-effects analyses. Other category includes multiracial, multiethnic, east Asian women. White women were used as the reference for all comparisons. NICHD=Eunice Kennedy Shriver National Institute of Child Health and Human Development.

51 studies included only high-risk women, eight (16%) only included low-risk women, and the rest had mixed risk groups or did not specify. 20 (39%) of 51 studies provided perinatal outcomes between five racial and ethnic groups (White, Black, south Asian, Hispanic, and other), 25 (49%) studies included four racial and ethnic groups (White, Black, south Asian, and other), five (10%) studies[34,42,51,65,72] included three groups, and one (2%) study[62] allowed comparison of perinatal outcomes between two racial and ethnic groups. Six datasets[42,48,51,70,79,84] provided data for more than 50% of the four confounders (maternal age, parity, BMI, and maternal educational attainment), 23 studies[36,37,42,44–46,48–53,56–58,60,63,64,71,72,78,80,83] provided data on three confounders, 19 studies[34,35,38–40,54,55,59,61,62,65,68,69,73–75,77,81,82] provided data on two confounders, and three studies[47,66,67] provided data on one confounder. Neonatal deaths were reported in nine studies (from 2 051 844 pregnancies)[34,35,39,41,48,49,68,73,76] and stillbirths were reported in six studies (from 380 017 pregnancies)[44,49,61,67,68,76] studies.

All comparative cohort studies evaluated with the Newcastle Ottawa Scale assessing the outcome of neonatal death had an overall low risk of bias (appendix pp 21–22). All studies had a low risk of bias for study selection and a medium risk of bias for comparability of cohorts. Six of nine studies had a low risk of bias for outcome assessment of the cohorts and three had a medium risk of bias. When assessing stillbirth, five of six studies had an overall low risk of bias, and one study had a medium risk of bias (appendix pp 21–22). All studies had a low risk of bias for study selection and a medium risk of bias for comparability of cohorts. Three studies had a low risk of bias for outcome assessment of the cohorts, two had a medium risk of bias, and one had a high risk of bias.

Nine studies provided IPD on race and ethnicity and neonatal deaths (12 468 neonatal deaths, 2 051 844 pregnancies; 12 countries).[34,35,39,41,48,49,68,73,76] Compared with White women, a higher risk of neonatal

	Maternal age	BMI	Parity	Educational attainment	Comparison		White women		Odds ratio (95% CI)	Weight (%)
					Events (n)	Women (N)	Events (n)	Women (N)		
Black women										
Goetzinger et al (2010)[67]	√				7	397	18	3282	2·40 (0·93–6·20)	13·01
Caritis et al (1998)[76]	√	√	√	√	40	1772	19	1077	1·28 (0·71–2·29)	24·99
Holzman et al (2001)[44]	√	√	√		9	743	5	2018	4·59 (1·42–14·81)	9·23
NICHD (2018)[49]	√	√	√		386	52518	370	116000	2·36 (2·04–2·74)	52·77
Overall					442	55430	412	122377	2·16 (1·46–3·19)	100·00
I^2=43·9%, p=0·148										
South Asian women										
Wright et al (2013)[61]	√		√		38	5532	16	4547	2·06 (1·14–3·71)	30·75
Huang et al (2010)[68]	√		√		19	35293	66	89350	0·75 (0·45–1·25)	33·23
NICHD (2018)[49]	√	√	√		22	9487	370	116000	0·75 (0·49–1·14)	36·01
Overall					79	50312	452	209897	1·02 (0·56–1·87)	100·00
I^2=76·9%, p=0·013										
Hispanic women										
Caritis et al (1998)[76]	√	√	√	√	5	388	19	1077	0·80 (0·27–2·38)	22·38
NICHD (2018)[49]	√	√	√		213	40409	370	116000	1·68 (1·42–1·98)	77·62
Overall					218	40797	389	117077	1·42 (0·78–2·60)	100·00
I^2=42·1%, p=0·189										
Other										
NICHD (2018)[49]	√	√	√		43	5512	370	116000	2·35 (1·72–3·23)	100·00
Overall									2·35 (1·72–3·23)	100·00

√ Adjusted for that confounder

0·07   1·00   15·00

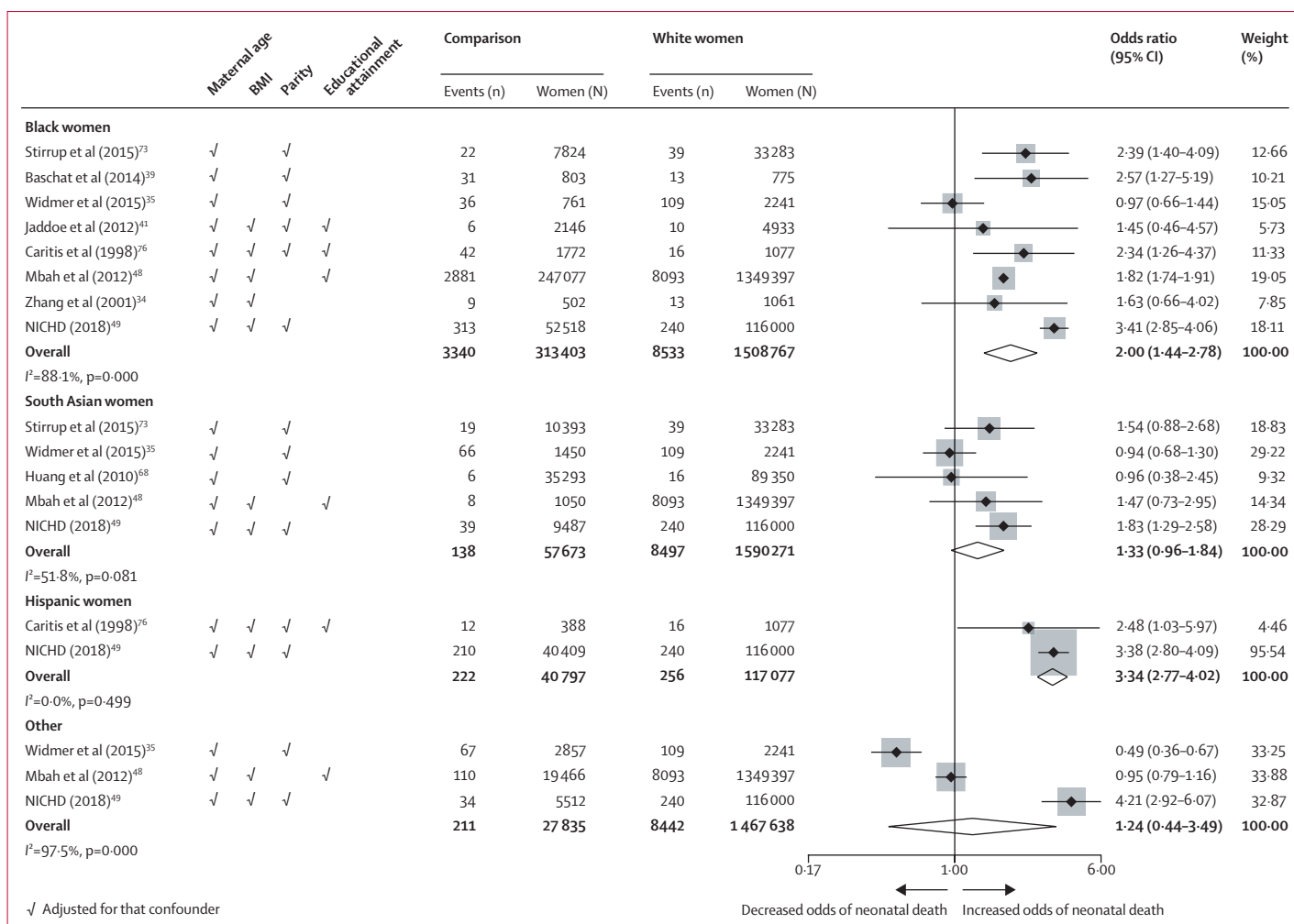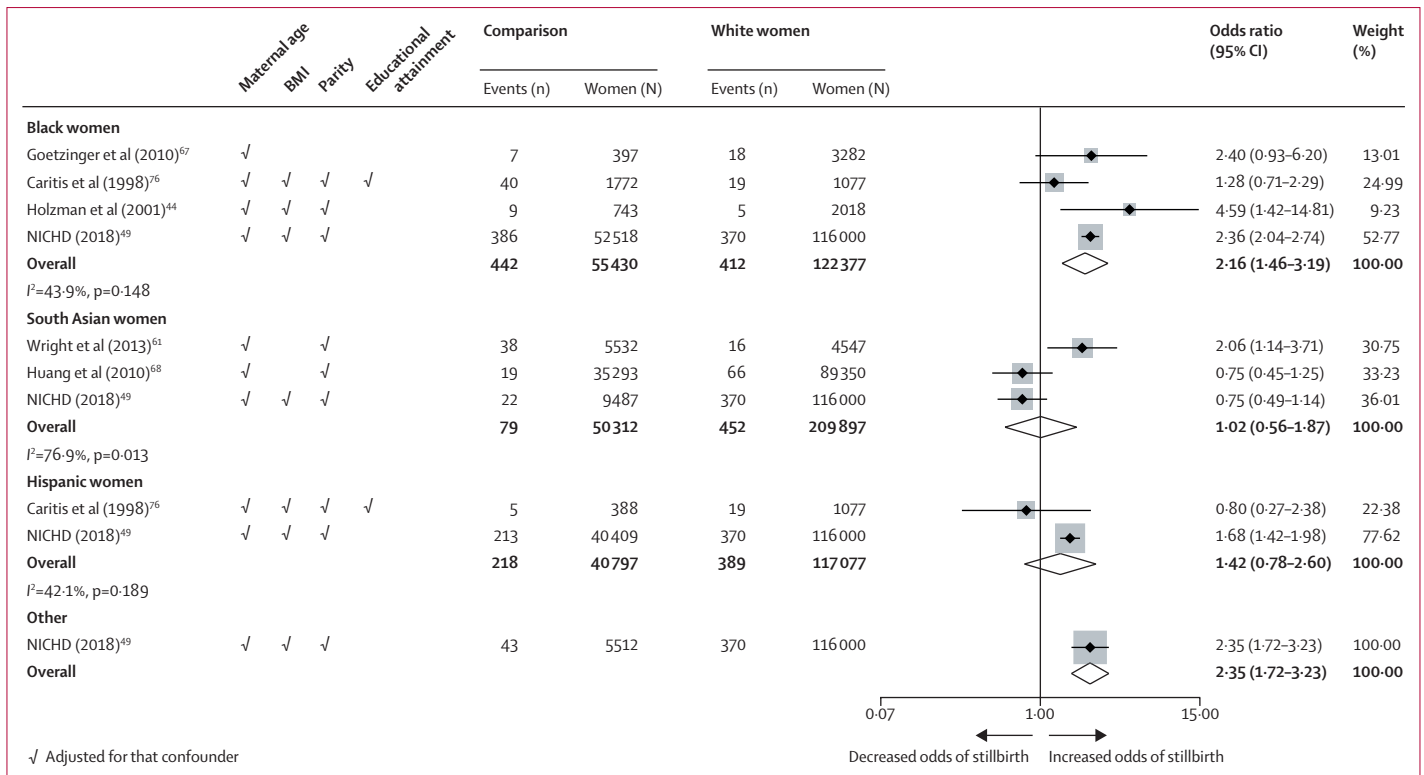Decreased odds of stillbirth   Increased odds of stillbirth

Figure 4: **Effect of race and ethnicity on the risk of stillbirths**
Weights are from random-effects analyses. Other category includes multiracial, multiethnic, east Asian women. White women were used as the reference for all comparisons. NICHD=Eunice Kennedy Shriver National Institute of Child Health and Human Development.

death was seen in Black women (OR 2·00, 95% CI 1·44–2·78) and Hispanic women (3·34, 2·77–4·02; figure 3). No differences were observed for south Asian women, when compared to White women, for neonatal death (OR 1·33, 95% CI 0·96–1·84).

Six studies provided IPD on race and ethnicity and stillbirths (1292 stillbirths, 380017 pregnancies; three countries).[44,49,61,67,68,76] The odds of stillbirth were two times higher in Black women (OR 2·16; 95% CI 1·46–3·19) and in other racial and ethnic groups (2·35, 1·72–3·23) when compared with White women (figure 4). No differences were observed in the odds of stillbirth in south Asian (OR 1·02, 95% CI 0·56–1·87) or Hispanic women (1·42, 0·78–2·60) when compared with White women. Subgroup analysis of neonatal death and stillbirth by geographical region was not possible because of the small number of studies.

45 studies provided IPD on race and ethnicity, and preterm birth (241817 preterm births, 2048987 pregnancies; 21 countries).[34–53,55–60,63–67,69,71–76,78–84] Compared with White women, a higher risk of preterm birth was seen in Black women (OR 1·65, 95% CI 1·46–1·88) and south Asian women (1·26, 1·07–1·48; appendix pp 23–25). 32 studies provided IPD on race and ethnicity and small for gestational age (266302 small-for-gestational-age babies, 1915004 pregnancies; 19 countries).[36–58,61,70,75–79] A higher risk of small-for-gestational-age babies was seen in

Black women (OR 1·39, 95% CI 1·13–1·72) and south Asian women (1·61, 1·32–1·95), when compared to White women (appendix pp 26–28). Subgroup analyses did not show significant variations in the effects of race and ethnicity on preterm births and small-for-gestational-age babies between regions (figures 5,6). Babies born to Black women were more likely to be preterm than those born to White women in the USA and Canada (OR 1·74, 95% CI 1·49–2·03), the UK (1·68, 1·23–2·31), and northern, western and southern Europe (OR 1·89, 95% CI 1·36–2·62) without any subgroup effect (p=0·41; figure 5A). The risks of preterm birth were high for south Asian women versus White women, with no variations between regions (p=0·30; figure 5B). The odds of small-for-gestational-age babies were increased for Black and south Asian women versus White women across regions, with no difference in the risk estimates between regions (p=0·27 for Black women and p=0·66 for south Asian women; figure 6).

Sensitivity analyses for the various assumptions of missingness used to impute data were consistent with the main findings for the association between race and ethnicity and perinatal outcomes (appendix pp 29–30). We did not find any collinearity issues among the covariates in the models, with the variance inflation factor below 5 (ranging between 1·00 and 2·12) in all models. Findings were similar when the analyses were limited to high-risk

pregnancies for all outcomes, except for increased odds of neonatal death in south Asian women versus White women (OR 1·85, 95% CI 1·05–3·26) and stillbirths in Hispanic versus White women (1·91, 1·45–2·51; appendix pp 31–34). The increased risk of small-for-gestational-age babies observed in south Asian versus White women was no longer present. Sensitivity analyses for complete cases (data not shown), by time of recruitment (appendix p 35), and after exclusion of one study[34] that was an outlier for recruitment period (1959–65; appendix p 36) showed findings similar to the main analysis for the outcomes. When one multi-country study that included women from low-income and middle-income countries was excluded,[35] the findings were similar to the main analysis for all outcomes, except for the risk of neonatal death, which became significant for south Asian versus White women (OR 1·62, 95% CI 1·25–2·10; appendix p 36).

Discussion

In high-income and middle-income countries, women from underserved and under-represented racial and ethnic groups are at increased risk of adverse perinatal outcomes. Black women are consistently at higher risk of all complications such as neonatal death, stillbirth, preterm birth, and small-for-gestational-age babies than White women. The effect varied for other racial and ethnic groups. Adverse outcomes such as preterm birth and small-for-gestational-age babies were higher in Black and south Asian women than in White women irrespective of the geographical region, and over time. Our work highlights the magnitude of disparities facing pregnant women from underserved racial and ethnic backgrounds irrespective of geographical region, emphasising the need for a broad global outlook to tackle these problems. To the best of our knowledge, our IPD meta-analysis is the largest and most comprehensive assessment to date of the magnitude of the association between race and ethnicity and adverse perinatal outcomes across high-income and upper-middle-income countries. Our work was based on a prospectively registered protocol with predefined aims and objectives. The harmonised IPPIC IPD data from multi-country cohorts provided us with a large sample size, facilitating

A	Black women		White women			Odds ratio (95% CI)	Weight (%)
	Event (n)	Women (N)	Event (n)	Women (N)			
USA and Canada							
Baschat et al (2014)[39]	155	803	86	775		1·96 (1·44–2·65)	4·12
Audibert et al (2010)[38]	8	62	40	720		2·45 (1·09–5·51)	1·71
Goetzinger et al (2010)[67]	38	397	279	3282		1·72 (1·24–2·39)	3·97
Caritis et al (1998)[76]	800	1772	542	1077		0·92 (0·78–1·08)	4·93
Sibai et al (1993)[79]	193	1551	45	570		1·61 (1·12–2·31)	3·76
Wright et al (2017)[58]	5	29	48	593		2·24 (0·81–6·19)	1·22
Holzman et al (2001)[44]	110	743	197	2018		1·63 (1·25–2·13)	4·36
Odibo et al (2011)[50]	62	325	77	735		1·45 (0·95–2·21)	3·40
Mbah et al (2012)[48]	50968	247077	139498	1349397		2·20 (2·17–2·22)	5·36
Zhang et al (2001)[34]	145	502	115	1061		2·94 (2·20–3·93)	4·21
NICHD (2018)[49]	9448	52518	13568	116000		1·70 (1·65–1·75)	5·35
Subtotal (I²=97·2%, p=0·000)						1·74 (1·49–2·03)	42·40
UK							
Allen et al (2017)[59]	12	108	16	398		2·75 (1·24–6·09)	1·75
Fraser et al (2013)[60]	14	131	677	12075		3·38 (0·78–14·60)	0·67
Chappell et al (1999)[80]	22	91	16	215		3·55 (1·65–7·63)	1·84
Carter et al (2017)[64]	52	193	165	881		1·38 (0·94–2·01)	3·66
Poston et al (2015)[83]	33	395	52	954		1·77 (1·11–2·82)	3·14
Stirrup et al (2015)[73]	726	7824	1707	33283		1·95 (1·78–2·14)	5·22
Velauthar et al (2014)[74]	6	90	21	401		1·23 (0·48–3·16)	1·37
H Al Wattar et al (2019)[43]	43	467	81	1403		1·54 (1·03–2·28)	3·55
Coomarasamy et al (2019)[81]	144	480	1014	2929		0·90 (0·72–1·11)	4·66
Subtotal (I²=84·1%, p=0·000)						1·68 (1·23–2·31)	25·86
Northern, western, and southern Europe							
Antsaklis et al (2000)[37]	12	49	374	3229		2·34 (1·21–4·53)	2·21
Llurba et al (2009)[47]	16	125	694	10415		2·14 (1·26–3·65)	2·79
Verlohren et al (2010)[56]	10	18	69	175		2·12 (0·75–6·01)	1·18
Jaddoe et al (2012)[41]	141	2146	296	4933		1·13 (0·90–1·41)	4·62
Prefumo et al (2008)[51]	5	41	13	203		3·79 (0·95–15·07)	0·74
Verlohren et al (2012)[57]	7	15	165	527		2·11 (0·74–6·01)	1·17
Lecarpentier et al (2013)[46]	47	125	17	93		3·16 (1·57–6·37)	2·06
Carbillon et al (2014)[63]	1532	19988	1105	14041		0·98 (0·90–1·06)	5·25
Langenveld et al (2011)[69]	7	11	43	177		5·14 (1·44–18·33)	0·86
van Oostwaard et al (2014)[55]	11	119	8	360		6·32 (2·39–16·71)	1·31
Crovetto et al (2016)[40]	13	218	397	6191		0·96 (0·54–1·69)	2·60
Subtotal (I²=80·4%, p=0·000)						1·89 (1·36–2·62)	24·78
Other							
Souza et al (2019)[52]	20	117	46	462		1·90 (1·06–3·41)	2·54
Subtotal						1·90 (1·06–3·41)	2·54
Multi-country							
Widmer et al (2015)[35]	85	761	321	2241		0·80 (0·62–1·03)	4·42
Subtotal						0·80 (0·62–1·03)	4·42
Overall (I²=96·3%, p=0·000)	64890	339291	161762	1571814		1·65 (1·46–1·88)	100·00

0·05   1·00   18·00

← Decreased odds of preterm birth   Increased odds of preterm birth →

(Figure 5 continues on next page)

B

	South Asian women		White women		Odds ratio (95% CI)	Weight (%)
	Event (n)	Women (N)	Event (n)	Women (N)		
USA and Canada						
Baschat et al (2014)[39]	10	88	86	775	1·01 (0·50–2·02)	2·83
Goetzinger et al (2010)[67]	13	112	279	3282	1·47 (0·84–2·59)	3·40
Holzman et al (2001)[64]	5	57	197	2018	0·87 (0·34–2·21)	2·02
Odibo et al (2011)[50]	6	94	77	735	0·72 (0·30–1·72)	2·21
Mbah et al (2012)[48]	163	1050	139 498	1 349 397	1·66 (1·40–1·96)	5·34
NICHD (2018)[49]	923	9487	13 568	116 000	0·76 (0·71–0·82)	5·59
Subtotal (I^2=93·2%, p=0·000)					1·06 (0·67–1·68)	21·39
UK						
Allen et al (2017)[59]	29	495	16	398	1·39 (0·74–2·63)	3·08
Fraser et al (2013)[60]	11	113	677	12 075	2·70 (0·65–11·20)	1·09
Carter et al (2017)[64]	18	85	165	881	1·04 (0·60–1·81)	3·47
Poston et al (2015)[83]	5	97	52	954	1·10 (0·43–2·86)	1·97
Stirrup et al (2015)[73]	779	10 393	1707	33 283	1·53 (1·40–1·68)	5·56
Velauthar et al (2014)[74]	43	568	21	401	1·50 (0·87–2·61)	3·48
Sovio et al (2015)[53]	7	91	195	3900	1·72 (0·78–3·78)	2·47
Dhillon-Smith et al (2019)[78]	50	102	163	403	1·41 (0·91–2·21)	4·00
H Al Wattar et al (2019)[43]	101	1206	81	1403	1·48 (1·09–2·02)	4·72
Coomarasamy et al (2019)[81]	113	282	1014	2929	1·28 (1·00–1·65)	4·99
Subtotal (I^2=0·0%, p=0·864)					1·49 (1·38–1·61)	34·85
Northern, western, and southern Europe						
Antsaklis et al (2000)[37]	6	32	374	3229	1·99 (0·80–4·91)	2·10
Jaddoe et al (2012)[41]	40	496	296	4933	1·39 (0·98–1·96)	4·54
Verlohren et al (2012)[57]	8	12	165	527	3·52 (1·04–11·97)	1·39
Staff et al (2005)[72]	6	13	57	222	2·57 (0·82–8·07)	1·53
Jenum et al (2010)[45]	11	200	23	379	1·46 (0·65–3·26)	2·43
Carbillon et al (2014)[63]	170	2404	1105	14 041	0·89 (0·75–1·05)	5·34
Crovetto et al (2016)[40]	32	635	397	6191	0·83 (0·57–1·21)	4·38
Subtotal (I^2=62·5%, p=0·014)					1·25 (0·91–1·71)	21·70
Other						
Makrides et al (2010)[82]	18	191	128	2111	1·82 (1·07–3·09)	3·58
Al-Amin et al (2018)[36]	12	96	40	431	1·35 (0·67–2·72)	2·81
Gurgel Alves et al (2014)[42]	53	556	7	128	1·63 (0·72–3·67)	2·40
Souza et al (2019)[52]	59	579	46	462	1·05 (0·69–1·61)	4·12
Subtotal (I^2=0·0%, p=0·431)					1·35 (1·02–1·79)	12·92
Multi-country						
North et al (2011)[71]	21	304	336	5061	1·04 (0·66–1·65)	3·95
Widmer et al (2015)[35]	159	1450	321	2241	0·85 (0·69–1·04)	5·20
Subtotal (I^2=0·0%, p=0·417)					0·88 (0·73–1·06)	9·15
Overall (I^2=86·8%, p=0·000)	2871	29 128	161 091	1 568 790	1·26 (1·07–1·48)	100·00

0·06   1·00   17·00

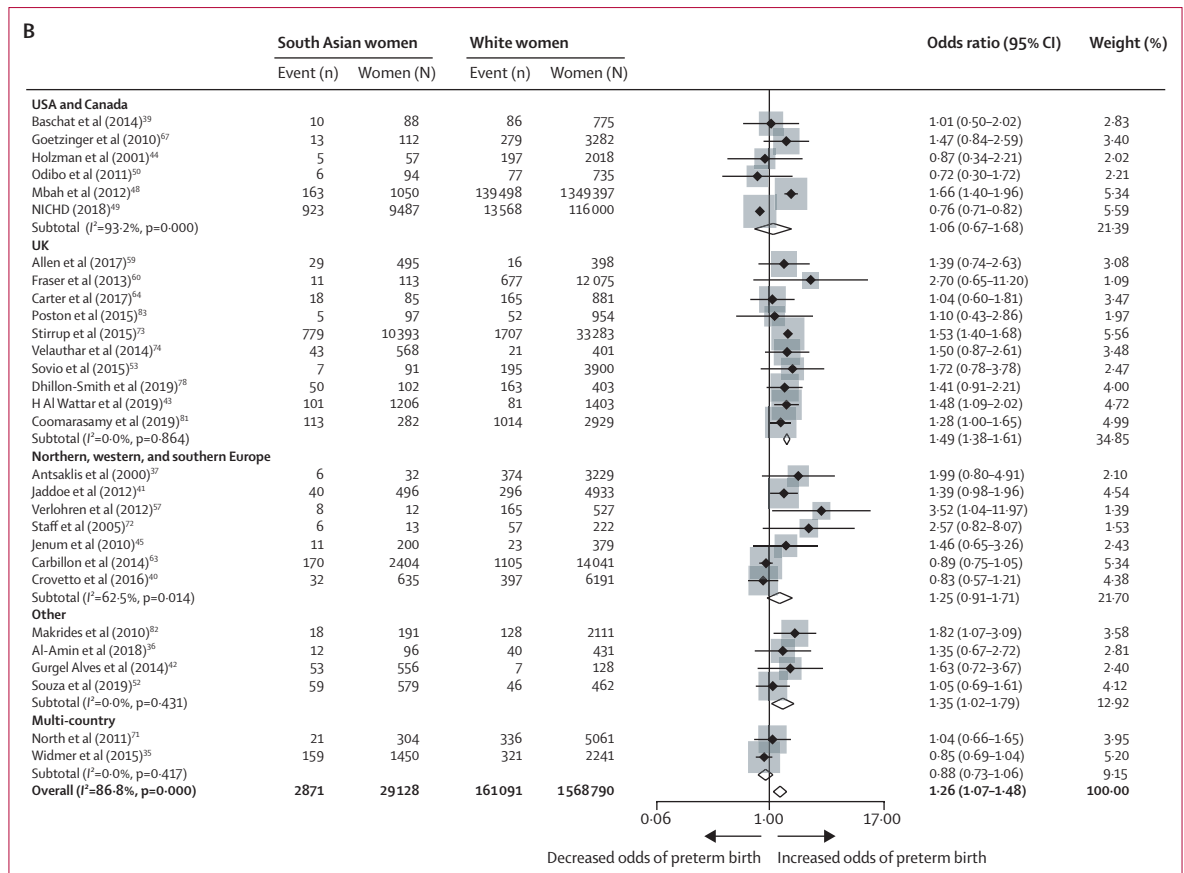← Decreased odds of preterm birth   Increased odds of preterm birth →

Figure 5: **Subgroup analysis of the effect of race and ethnicity on the risk of preterm births by region**
Weights are from random-effects analysis. Adjusted for maternal age, BMI, parity, and maternal educational attainment. Other category includes studies from Australia and Brazil. (A) Effects of race and ethnicity on preterm births for Black women. Subgroup effect: p=0·408. (B) Effects of race and ethnicity on preterm births for south Asian women. Subgroup effect: p=0·296. White women were the reference group in all comparisons. NICHD=Eunice Kennedy Shriver National Institute of Child Health and Human Development.

high precision in the findings and increasing the generalisability of these results. We used multiple imputations to deal with missing variables, thereby avoiding the loss of useful information.

We considered race and ethnicity to be social constructs without biological meaning and reported these terms in line with current recommendations to minimise bias.[19] We used the terms race and ethnicity as a lens through which to study the disparities in pregnancy outcomes in women from underserved and under-represented groups because of differential treatment and access to health care. We considered the effects of race and ethnicity to include the effects of a woman's appearance (phenotype including skin colour) that influences how she is perceived by others, and also the understanding of her appearance affecting her identity and behaviour, her parents' appearances, and the cultural context.[11] Given the different ways in which women of African origin might self-identify their origin, and the varied reporting of these women, we categorised the grouping as Black for the purpose of our analysis, as recommended by current guidance.[19] Our subgroup analysis assessed the variations

in outcome disparities in Black women by geographical region. The term Asian is broad and includes numerous countries of origin (eg, Bangladesh, China, India, Indonesia, Japan, and Pakistan). Given the significant differences between the various Asian ethnic groups in terms of rates of diabetes, hypertension, and other adverse outcomes, instead of pooling in one category, we reported them separately as south Asians and east Asians.[85] Because of the small sample size of east Asians in the IPPIC dataset, we included them in the "other" group for the purpose of analysis. We classified all women of Hispanic identity under the underserved and under-represented race and ethnic category, including those who might identify as White Hispanic. We did so on the basis of how women might be perceived by others, which can affect their experiences and expose them to inequalities in care. In a survey, Hispanic adults said that they are described by most people as Hispanic rather than White.[20] In our study, we considered White women to be the reference group through the lens of societal context, irrespective of their majority or minority status,[18] where White experience is one of privilege and power across
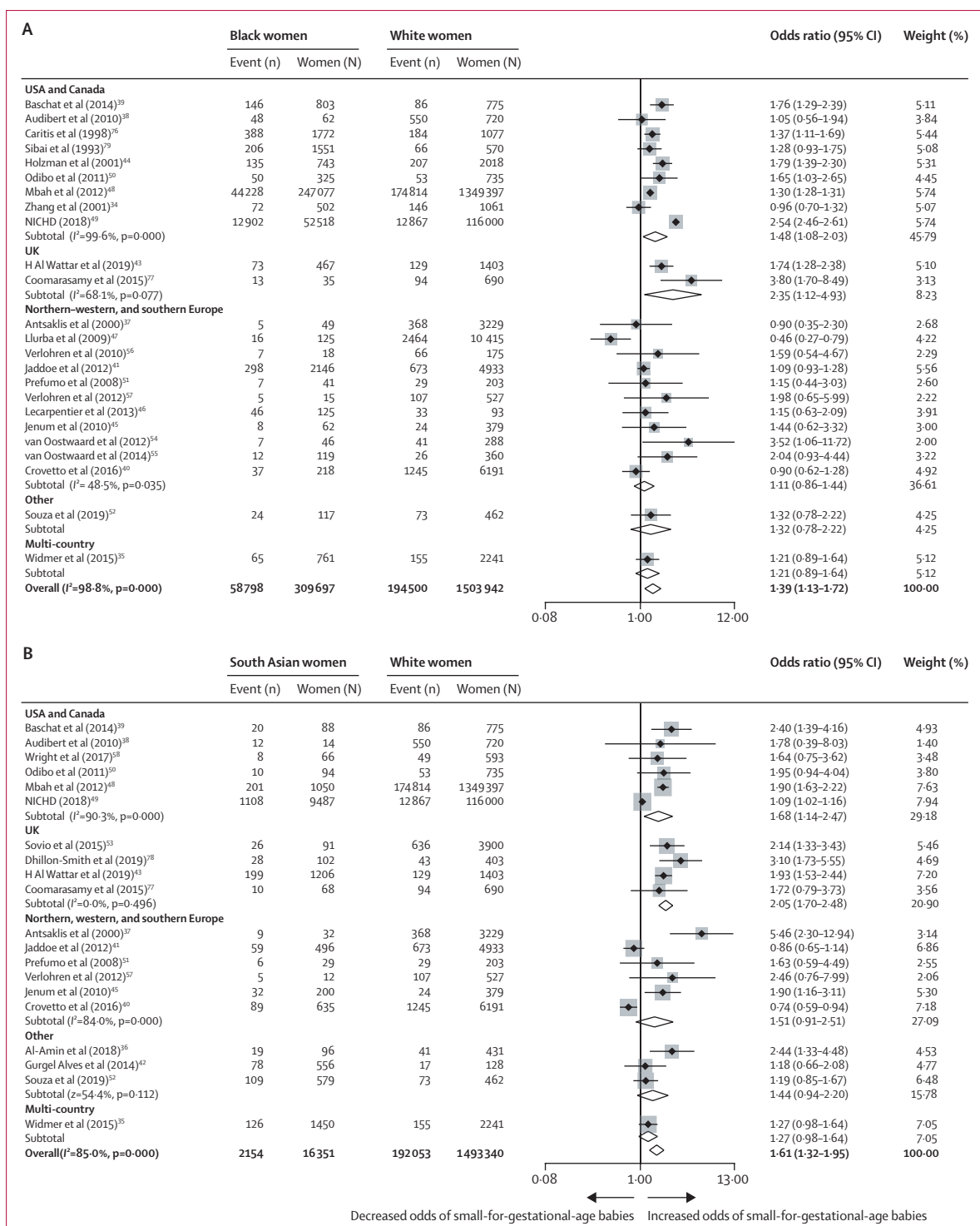
A

	Black women		White women			Odds ratio (95% CI)	Weight (%)
	Event (n)	Women (N)	Event (n)	Women (N)			
USA and Canada							
Baschat et al (2014)[39]	146	803	86	775		1·76 (1·29–2·39)	5·11
Audibert et al (2010)[38]	48	62	550	720		1·05 (0·56–1·94)	3·84
Caritis et al (1998)[76]	388	1772	184	1077		1·37 (1·11–1·69)	5·44
Sibai et al (1993)[79]	206	1551	66	570		1·28 (0·93–1·75)	5·08
Holzman et al (2001)[44]	135	743	207	2018		1·79 (1·39–2·30)	5·31
Odibo et al (2011)[50]	50	325	53	735		1·65 (1·03–2·65)	4·45
Mbah et al (2012)[48]	44228	247077	174814	1349397		1·30 (1·28–1·31)	5·74
Zhang et al (2001)[34]	72	502	146	1061		0·96 (0·70–1·32)	5·07
NICHD (2018)[49]	12902	52518	12867	116000		2·54 (2·46–2·61)	5·74
Subtotal (I²=99·6%, p=0·000)						1·48 (1·08–2·03)	45·79
UK							
H Al Wattar et al (2019)[43]	73	467	129	1403		1·74 (1·28–2·38)	5·10
Coomarasamy et al (2015)[77]	13	35	94	690		3·80 (1·70–8·49)	3·13
Subtotal (I²=68·1%, p=0·077)						2·35 (1·12–4·93)	8·23
Northern–western, and southern Europe							
Antsaklis et al (2000)[37]	5	49	368	3229		0·90 (0·35–2·30)	2·68
Llurba et al (2009)[47]	16	125	2464	10415		0·46 (0·27–0·79)	4·22
Verlohren et al (2010)[56]	7	18	66	175		1·59 (0·54–4·67)	2·29
Jaddoe et al (2012)[41]	298	2146	673	4933		1·09 (0·93–1·28)	5·56
Prefumo et al (2008)[51]	7	41	29	203		1·15 (0·44–3·03)	2·60
Verlohren et al (2012)[57]	5	15	107	527		1·98 (0·65–5·99)	2·22
Lecarpentier et al (2013)[46]	46	125	33	93		1·15 (0·63–2·09)	3·91
Jenum et al (2010)[45]	8	62	24	379		1·44 (0·62–3·32)	3·00
van Oostwaard et al (2012)[54]	7	46	41	288		3·52 (1·06–11·72)	2·00
van Oostwaard et al (2014)[55]	12	119	26	360		2·04 (0·93–4·44)	3·22
Crovetto et al (2016)[40]	37	218	1245	6191		0·90 (0·62–1·28)	4·92
Subtotal (I²=48·5%, p=0·035)						1·11 (0·86–1·44)	36·61
Other							
Souza et al (2019)[52]	24	117	73	462		1·32 (0·78–2·22)	4·25
Subtotal						1·32 (0·78–2·22)	4·25
Multi-country							
Widmer et al (2015)[35]	65	761	155	2241		1·21 (0·89–1·64)	5·12
Subtotal						1·21 (0·89–1·64)	5·12
Overall (I²=98·8%, p=0·000)	58798	309697	194500	1503942		1·39 (1·13–1·72)	100·00

0·08   1·00   12·00

B

	South Asian women		White women			Odds ratio (95% CI)	Weight (%)
	Event (n)	Women (N)	Event (n)	Women (N)			
USA and Canada							
Baschat et al (2014)[39]	20	88	86	775		2·40 (1·39–4·16)	4·93
Audibert et al (2010)[38]	12	14	550	720		1·78 (0·39–8·03)	1·40
Wright et al (2017)[58]	8	66	49	593		1·64 (0·75–3·62)	3·48
Odibo et al (2011)[50]	10	94	53	735		1·95 (0·94–4·04)	3·80
Mbah et al (2012)[48]	201	1050	174814	1349397		1·90 (1·63–2·22)	7·63
NICHD (2018)[49]	1108	9487	12867	116000		1·09 (1·02–1·16)	7·94
Subtotal (I²=90·3%, p=0·000)						1·68 (1·14–2·47)	29·18
UK							
Sovio et al (2015)[53]	26	91	636	3900		2·14 (1·33–3·43)	5·46
Dhillon-Smith et al (2019)[78]	28	102	43	403		3·10 (1·73–5·55)	4·69
H Al Wattar et al (2019)[43]	199	1206	129	1403		1·93 (1·53–2·44)	7·20
Coomarasamy et al (2015)[77]	10	68	94	690		1·72 (0·79–3·73)	3·56
Subtotal (I²=0·0%, p=0·496)						2·05 (1·70–2·48)	20·90
Northern, western, and southern Europe							
Antsaklis et al (2000)[37]	9	32	368	3229		5·46 (2·30–12·94)	3·14
Jaddoe et al (2012)[41]	59	496	673	4933		0·86 (0·65–1·14)	6·86
Prefumo et al (2008)[51]	6	29	29	203		1·63 (0·59–4·49)	2·55
Verlohren et al (2012)[57]	5	12	107	527		2·46 (0·76–7·99)	2·06
Jenum et al (2010)[45]	32	200	24	379		1·90 (1·16–3·11)	5·30
Crovetto et al (2016)[40]	89	635	1245	6191		0·74 (0·59–0·94)	7·18
Subtotal (I²=84·0%, p=0·000)						1·51 (0·91–2·51)	27·09
Other							
Al-Amin et al (2018)[36]	19	96	41	431		2·44 (1·33–4·48)	4·53
Gurgel Alves et al (2014)[42]	78	556	17	128		1·18 (0·66–2·08)	4·77
Souza et al (2019)[52]	109	579	73	462		1·19 (0·85–1·67)	6·48
Subtotal (z=54·4%, p=0·112)						1·44 (0·94–2·20)	15·78
Multi-country							
Widmer et al (2015)[35]	126	1450	155	2241		1·27 (0·98–1·64)	7·05
Subtotal						1·27 (0·98–1·64)	7·05
Overall (I²=85·0%, p=0·000)	2154	16351	192053	1493340		1·61 (1·32–1·95)	100·00

0·08   1·00   13·00

◄──── Decreased odds of small-for-gestational-age babies   Increased odds of small-for-gestational-age babies ────►

Figure 6: Subgroup analysis of the effect of race and ethnicity on the risk of small-for-gestational-age babies by region
Weights are from random-effects analysis. Adjusted for maternal age, BMI, parity, and maternal educational attainment. Other category includes studies from Australia and Brazil. (A) Effects of race and ethnicity on small-for-gestational-age babies for Black women. Subgroup effect: p=0·267. (B) Effects of race and ethnicity on small-for-gestational-age babies for south Asian women. Subgroup effect: p=0·661. White women were the reference group in all comparisons. NICHD=Eunice Kennedy Shriver National Institute of Child Health and Human Development.

regions and settings[86] and White women are expected to have optimal outcomes compared with other groups.

Confounding variables adjusted for in our analysis were identified a priori by use of a directed acyclic graph, and unlike previous studies in this area[87–90] we refrained from the unnecessary adjustment of gestational age and birthweight due to their collider status.[30,91] By adjusting for the highest educational level attained as a measure of socioeconomic status, we avoided overadjusting for other factors along the pathway.[92] Since the highest educational attainment achieved by an individual is usually reached in early adulthood and is the main marker for upward mobility,[93] we consider it to be a key marker of social status such as income, employment, and living environment.[93–96] Studies show that the association between education and health is driven by increases in human capital, with people who have lower levels of education experiencing a faster health decline than those with higher levels of education.[97]

Our study had some limitations. We only included cohorts of pregnant women shared and harmonised as part of the IPPIC project, and data from studies not in the IPPIC data repository were not considered in the analysis. There were high levels of missing data in variables in some of the cohorts used for the IPD meta-analysis. However, our sensitivity analysis on complete cases resulted in similar results to our imputed dataset. Some of the cohorts included pregnant women over many decades, and the risk of adverse perinatal outcomes could have changed over time. Stillbirth was also variably defined within individual cohorts, which might have affected estimates in our analysis. Our analysis did not consider unmeasured factors that could confound the association between race and ethnicity and perinatal outcomes. The definitions of race and ethnicity differed between studies according to the databases used, the geographical regions, and time of data collection within the included IPPIC cohorts. We were only able to assess for variations in disparities due to race and ethnicity in perinatal outcomes between geographical regions, and not by health systems (private sector, government funded, or mixed models) because of the paucity of reported data.

We acknowledge that the experiences and challenges faced by women from ethnic groups such as south Asians might vary between regions due to differences in historical immigration patterns (eg, migration to escape civil war, for economic reasons, or to join family members) and policies.[98] But overall, we did not observe significant differences between the subgroups categorised by region for increased risk of small-for-gestational-age babies in women of south Asian ethnicity. It is likely that the effects of race and ethnicity on perinatal outcomes might be different within subgroups such as African and African Caribbean women born in a high-income country (eg, the UK, USA, or Canada) compared to first-generation migrants to that country, and also between various south Asian groups (eg, those of Bangladeshi, Indian, and Pakistani origin) who have been reported to have varied health outcomes, such as the highest levels of infant mortality rates in babies born to women of Pakistani origin.[99–101] However, we were limited by the paucity of relevant data in the primary studies and were not able to undertake this analysis. We were able to adjust for only one measure of a woman's socioeconomic status, maternal educational attainment, and not for other measures such as income and occupation, because of the availability of sparse and heterogeneous data in the IPPIC repository. Since the studies involved in our meta-analyses were not specifically done to assess the effects of race and ethnicity on perinatal outcomes, it is difficult to interpret the likelihood of the publication of a study included in our IPD meta-analysis with the magnitude of the association we estimated or the precision of these estimates. Therefore, we refrained from assessing the risk of publication bias.[11]

Since the 1980s, neonatal mortality rates have been on the decline in most countries, but this overall trend hides underlying differences within individual racial and ethnic groups.[102] For example, in the UK, a 12% fall in stillbirths among White women between 2013 and 2018 contrasts starkly with a contemporaneous 5% rise in stillbirths among Black women.[1] The effect of race and ethnicity has often been shown to be associated with adverse perinatal outcomes, but this has mostly been presented in the light of it being modified by socioeconomic status.[1,9] Studies such as the UK National Maternity and Perinatal Audit[9] and those from the USA[103,104] report higher rates of adverse perinatal outcomes in Black and Asian women, as well as women from other underserved groups, than in White women even after adjusting for socioeconomic deprivation,[9] implying the contribution of other factors.[105]

Our study shows that after controlling for the effect of maternal characteristics, including a woman's educational attainment, the association between race and ethnicity and adverse perinatal outcomes persists. Complex multifactorial characteristics influence these outcomes in women from underserved racial and ethnic groups. The unique set of challenges posed by pregnancy is further worsened in individuals who are disadvantaged by their sex, race, and ethnicity.[106] Racial discrimination is known to be associated with chronic stress that can influence pregnancy outcomes.[4] Furthermore, women from underserved racial and ethnic groups encounter discrimination at various levels, contributing to adverse pregnancy outcomes: at the institutional level, leading to differential access to antenatal care; at the interpersonal level in their interactions with health-care professionals who do not acknowledge their concerns; and through internalised racism, where women from marginalised groups accept their perceived incompetence that limits them from seeking timely care.[4] These problems are compounded by racial discrimination across generations and gaps in health literacy,[4] which are in turn affected by the environment, social relationships, and employment

opportunities.[5] Previous studies have incorrectly adjusted for birthweight and gestational age at delivery, which dampens the true effect of race and ethnicity on adverse perinatal outcomes.[9]

Our finding of disparities in perinatal outcomes across regions and over time in underserved racial and ethnic groups highlights the global need to address the structural, interpersonal, and internalised barriers faced by these women. In many countries, poor maternal and perinatal outcomes have been linked to structural racism[7,107–110]—a system where public policies, institutional practices, cultural depiction, and other means contribute to and reinforce racial inequity.[111] The recent inquiry into racial justice and human rights in UK maternity care found that systemic factors such as negative stereotyping, microaggressions, race-based risk assumptions, and dehumanisation of women from underserved racial and ethnic groups contributed to their poor pregnancy outcomes.[107] Structural racism was also highlighted as a key contributing factor to poor outcomes in Black mothers in the testimonies submitted to the US House Oversight and Reform Committee for their hearing, Birthing While Black.[112]

Multifaceted antenatal interventions are urgently needed across all regions and countries to reduce the racial and ethnic inequities in pregnancy care and outcomes. Central to any such effort should be the removal of organisational and policy-level structural barriers contributing to poor perinatal outcomes.[113] Interventions should focus on understanding why Black and south Asian babies die or develop complications at a disproportionate rate to White babies, and avoid clinical decisions guided by race and ethnicity that could exacerbate such inequalities.[114] In the UK, the Royal College of Obstetricians and Gynaecologists has launched the Race Equality Taskforce to tackle racial disparities in women's health care, including pregnancy outcomes.[115] This is supported by national strategies such as the Race and Health Observatory, the NHS England Equality strategy, and the Core20Plus5 approach.[116–118] Similar efforts are underway in other countries.[119,120]

The window of opportunity available to maternity services to tackle these disparities is brief but substantial, and requires resource-intensive and time-consuming changes in social and maternity care.[105] These efforts need to be complemented by a life course approach to optimising the health of, and underpinning determinants in, girls and women from underserved and under-represented groups. The curriculum and training offered to midwifery and medical students should integrate strategies to identify explicit and implicit racial biases in health-care settings and provide the tools to improve communication while caring for women from various backgrounds.[121]

Despite race and ethnicity being a risk factor for adverse health outcomes, particularly in pregnancy, there are no comprehensive research strategies or initiatives to address this problem. In addition to encouraging women from various racial and ethnic backgrounds to participate in research,[122] funding bodies need to prioritise topics that directly address the disparities in pregnancy outcomes that are related to race and ethnicity. The voices of women from relevant backgrounds should be central to lead and guide the efforts in this area. Given that race and ethnicity are key demographic variables, studies should aim to comprehensively collect and report these data in line with current recommendations at all stages of a woman's life.[19] This will allow us to not only map the magnitude of disparities at various timepoints, such as childhood, adolescence, pre-pregnancy, and pregnancy, contributing to poor pregnancy outcomes and their long-term effects in later years, but also plan targeted interventions at crucial timepoints to improve the health of babies in the short and long term, with the impact spanning generations.

References
1 Draper ES, Gallimore ID, Smith LK, et al. MBRRACE-UK Perinatal Mortality Surveillance Report. UK Perinatal Deaths for Births from January to December 2018. December, 2020. https://www.npeu.ox.ac.uk/assets/downloads/mbrrace-uk/reports/maternal-report-2021/MBRRACE-UK_Maternal_Report_2021_-_FINAL_-_WEB_VERSION.pdf (accessed June 30, 2022).
2 Bryant AS, Worjoloh A, Caughey AB, Washington AE. Racial/ethnic disparities in obstetric outcomes and care: prevalence and determinants. *Am J Obstet Gynecol* 2010; **202:** 335–43.
3 Parchem JG, Rice MM, Grobman WA, et al. Racial and ethnic disparities in adverse perinatal outcomes at term. *Am J Perinatol* 2021; published online May 31. https://doi.org/10.1055/s-0041-1730348.
4 Alhusen JL, Bower KM, Epstein E, Sharps P. Racial discrimination and adverse birth outcomes: an integrative review. *J Midwifery Womens Health* 2016; **61:** 707–20.
5 Lu MC, Halfon N. Racial and ethnic disparities in birth outcomes: a life-course perspective. *Matern Child Health J* 2003; **7:** 13–30.
6 Wallace ME, Mendola P, Kim SS, et al. Racial/ethnic differences in preterm perinatal outcomes. *Am J Obstet Gynecol* 2017; **216:** 306.e1–12.

7   Taylor JK. Structural racism and maternal health among black women. *J Law Med Ethics* 2020; **48:** 506–17.

8   Braveman P, Dominguez TP, Burke W, et al. Explaining the Black–White disparity in preterm birth: a consensus statement from a multi-disciplinary scientific work group convened by the March of Dimes. *Front Reprod Heal* 2021; **3:** 684207.

9   Jardine J, Walker K, Gurol-Urganci I, et al. Adverse pregnancy outcomes attributable to socioeconomic and ethnic inequalities in England: a national cohort study. *Lancet* 2021; **398:** 1905–12.

10   Martin JA, Hamilton B, E Osterman MJK, Driscoll AK. Births: final data for 2018. *National Vital Statistics Reports.* Nov 27, 2019. https://stacks.cdc.gov/view/cdc/82909 (accessed June 30, 2022).

11   VanderWeele TJ, Robinson WR. On the causal interpretation of race in regressions adjusting for confounding and mediating variables. *Epidemiology* 2014; **25:** 473–84.

12   Sheikh J, Allotey J, Fernandez-Felix BM, Zamora J, Khalil A, Thangaratinam S. Effect of race on the risk of adverse perinatal outcomes: an individual participant data meta-analysis protocol. OSF Registries. Jan 14, 2022. https://doi.org/10.17605/OSF.IO/M3S2A (accessed June 30, 2022).

13   Stewart LA, Clarke M, Rovers M, et al. Preferred reporting items for a systematic review and meta-analysis of individual participant data. *JAMA* 2015; **313:** 1657.

14   Townsend R, Khalil A, Premakumar Y, et al. Prediction of pre-eclampsia: review of reviews. *Ultrasound Obstet Gynecol* 2019; **54:** 16–27.

15   Allotey J, Whittle R, Snell KIE, et al. External validation of prognostic models to predict stillbirth using the International Prediction of Pregnancy Complications (IPPIC) Network database: an individual participant data meta-analysis. *Ultrasound Obstet Gynecol* 2022; **59:** 209–19.

16   Allotey J, Snell KI, Smuk M, et al. Validation and development of models using clinical, biochemical and ultrasound markers for predicting pre-eclampsia: an individual participant data meta-analysis. *Health Technol Assess* 2020; **24:** 1–252.

17   The World Bank Group. Data: World Bank country and lending groups. 2022. https://datahelpdesk.worldbank.org/knowledgebase/articles/906519-world-bank-country-and-lending-groups (accessed June 30, 2022).

18   Jugert P, Kaiser MJ, Ialuna F, Civitillo S. Researching race-ethnicity in race-mute Europe. *Infant Child Dev* 2022; **31:** e2260.

19   Flanagin A, Frey T, Christiansen SL. Updated guidance on the reporting of race and ethnicity in medical and science journals. *JAMA* 2021; **326:** 621–27.

20   Noe-Bustamante L, Gonzalez-Barrera A, Edwards K, Mora L, Hugo Lopez M. Majority of Latinos say skin color impacts opportunity in America and shapes daily life: 4. Measuring the racial identity of Latinos. *Pew Research Center*. Nov 4, 2021. https://www.pewresearch.org/hispanic/2021/11/04/measuring-the-racial-identity-of-latinos/ (accessed June 30, 2022).

21   WHO. Neonatal and perinatal mortality: country, regional and global estimates. Geneva, World Health Organization, 2006.

22   US Centers for Disease Control and Prevention. CDC 24/7: saving lives, protecting people. What is stillbirth? Nov 16, 2020. https://www.cdc.gov/ncbddd/stillbirth/facts.html (accessed July 22, 2022).

23   WHO. Born too soon: the global action report on preterm birth. Geneva: World Health Organization, 2012.

24   Royal College of Obstetrics & Gynaecologists. Investigation and management of small-for-gestational-age fetus. Green-top guideline No. 31. February, 2013. https://www.rcog.org.uk/media/t3lmjhnl/gtg_31.pdf (accessed July 22, 2022).

25   Wells G, Shea B, O'Connell O, et al. The Newcastle-Ottawa Scale (NOS) for assessing the quality of nonrandomised studies in meta-analyses. Third symposium on systematic reviews beyond the basics: improving quality and impact. July 3–5, 2000; Oxford, UK. https://www.ohri.ca/programs/clinical_epidemiology/oxford.asp (accessed July 27, 2022).

26   Moons KGM, Donders RART, Stijnen T, Harrell FE Jr. Using the outcome for imputation of missing predictor values was preferred. *J Clin Epidemiol* 2006; **59:** 1092–101.

27   von Hippel PT. 4. Regression with missing Ys: an improved strategy for analyzing multiply imputed data. *Sociol Methodol* 2007; **37:** 83–117.

28   Belsley D, Kuh E, Welsch R. Detecting and assessing collinearity. In: Belsley D, Kuh E, Welsch R, eds. Regression diagnostic: identifying influential data and sources of collinearity. Hoboken, NJ: John Wiley & Sons, 1980: 85–191.

29   Greenland S, Pearl J, Robins JM. Causal diagrams for epidemiologic research. *Epidemiology* 1999; **10:** 37–48.

30   Wilcox AJ, Weinberg CR, Basso O. On the pitfalls of adjusting for gestational age at birth. *Am J Epidemiol* 2011; **174:** 1062–68.

31   Rubin DB. Multiple Imputation for nonresponse in surveys. Hoboken, NJ: John Wiley & Sons, 1987.

32   DerSimonian R, Laird N. Meta-analysis in clinical trials. *Control Clin Trials* 1986; **7:** 177–88.

33   UN Statistics Division. Methodology: standard country or area codes for statistical use (M49). United Nations. 2022. https://unstats.un.org/unsd/methodology/m49/ (accessed June 30, 2022).

34   Zhang J, Troendle JF, Levine RJ. Risks of hypertensive disorders in the second pregnancy. *Paediatr Perinat Epidemiol* 2001; **15:** 226–31.

35   Widmer M, Cuesta C, Khan KS, et al. Accuracy of angiogenic biomarkers at ≤20weeks' gestation in predicting the risk of pre-eclampsia: a WHO multicentre study. *Pregnancy Hypertens* 2015; **5:** 330–38.

36   Al-Amin A, Rolnik DL, Black C, et al. Accuracy of second trimester prediction of preterm preeclampsia by three different screening algorithms. *Aust N Z J Obstet Gynaecol* 2018; **58:** 192–96.

37   Antsaklis A, Daskalakis G, Tzortzis E, Michalas S. The effect of gestational age and placental location on the prediction of pre-eclampsia by uterine artery Doppler velocimetry in low-risk nulliparous women. *Ultrasound Obstet Gynecol* 2000; **16:** 635–39.

38   Audibert F, Boucoiran I, An N, et al. Screening for preeclampsia using first-trimester serum markers and uterine artery Doppler in nulliparous women. *Am J Obstet Gynecol* 2010; **203:** 383.e1–8.

39   Baschat AA, Magder LS, Doyle LE, Atlas RO, Jenkins CB, Blitzer MG. Prediction of preeclampsia utilizing the first trimester screening examination. *Am J Obstet Gynecol* 2014; **211:** 514.e1–7.

40   Crovetto F, Triunfo S, Crispi F, et al. First-trimester screening with specific algorithms for early- and late-onset fetal growth restriction. *Ultrasound Obstet Gynecol* 2016; **48:** 340–48.

41   Jaddoe VWV, van Duijn CM, Franco OH, et al. The Generation R Study: design and cohort update 2012. *Eur J Epidemiol* 2012; **27:** 739–56.

42   Gurgel Alves JA, Praciano de Sousa PC, Bezerra Maia E, Holanda Moura S, Kane SC, da Silva Costa F. First-trimester maternal ophthalmic artery Doppler analysis for prediction of pre-eclampsia. *Ultrasound Obstet Gynecol* 2014; **44:** 411–18.

43   H Al Wattar B, Dodds J, Placzek A, et al. Mediterranean-style diet in pregnant women with metabolic risk factors (ESTEEM): a pragmatic multicentre randomised trial. *PLoS Med* 2019; **16:** e1002857.

44   Holzman C, Bullen B, Fisher R, Paneth N, Reuss L. Pregnancy outcomes and community health: the POUCH study of preterm delivery. *Paediatr Perinat Epidemiol* 2001; **15** (suppl 2)**:** 136–58.

45   Jenum AK, Sletner L, Voldner N, et al. The STORK Groruddalen research programme: a population-based cohort study of gestational diabetes, physical activity, and obesity in pregnancy in a multiethnic population. Rationale, methods, study population, and participation rates. *Scand J Public Health* 2010; **38** (suppl)**:** 60–70.

46   Lecarpentier E, Tsatsaris V, Goffinet F, Cabrol D, Sibai B, Haddad B. Risk factors of superimposed preeclampsia in women with essential chronic hypertension treated before pregnancy. *PLoS One* 2013; **8:** e62140.

47   Llurba E, Carreras E, Gratacós E, et al. Maternal history and uterine artery Doppler in the assessment of risk for development of early- and late-onset preeclampsia and intrauterine growth restriction. *Obstet Gynecol Int* 2009; **2009:** 275613.

48   Mbah AK, Sharma PP, Alio AP, Fombo DW, Bruder K, Salihu HM. Previous cesarean section, gestational age at first delivery and subsequent risk of pre-eclampsia in obese mothers. *Arch Gynecol Obstet* 2012; **285:** 1375–81.

49   Eunice Kennedy Shriver National Institute of Child Health and Human Development. NICHD DASH Consortium of Safe Labor (CSL). 2018. https://dash.nichd.nih.gov/study/2331 (accessed June 30, 2022).

50 Odibo AO, Zhong Y, Goetzinger KR, et al. First-trimester placental protein 13, PAPP-A, uterine artery Doppler and maternal characteristics in the prediction of pre-eclampsia. *Placenta* 2011; **32:** 598–602.

51 Prefumo F, Fratelli N, Ganapathy R, Bhide A, Frusca T, Thilaganathan B. First trimester uterine artery Doppler in women with previous pre-eclampsia. *Acta Obstet Gynecol Scand* 2008; **87:** 1271–75.

52 Souza RT, Cecatti JG, Costa ML, et al. Planning, implementing, and running a multicentre preterm birth study with biobank resources in Brazil: the Preterm SAMBA Study. *BioMed Res Int* 2019; **2019:** 5476350.

53 Sovio U, White IR, Dacey A, Pasupathy D, Smith GCS. Screening for fetal growth restriction with universal third trimester ultrasonography in nulliparous women in the Pregnancy Outcome Prediction (POP) study: a prospective cohort study. *Lancet* 2015; **386:** 2089–97.

54 van Oostwaard MF, Langenveld J, Bijloo R, et al. Prediction of recurrence of hypertensive disorders of pregnancy between 34 and 37 weeks of gestation: a retrospective cohort study. *BJOG* 2012; **119:** 840–47.

55 van Oostwaard MF, Langenveld J, Schuit E, et al. Prediction of recurrence of hypertensive disorders of pregnancy in the term period, a retrospective cohort study. *Pregnancy Hypertens* 2014; **4:** 194–202.

56 Verlohren S, Galindo A, Schlembach D, et al. An automated method for the determination of the sFlt-1/PlGF ratio in the assessment of preeclampsia. *Am J Obstet Gynecol* 2010; **202:** 161.e1–11.

57 Verlohren S, Herraiz I, Lapaire O, et al. The sFlt-1/PlGF ratio in different types of hypertensive pregnancy disorders and its prognostic potential in preeclamptic patients. *Am J Obstet Gynecol* 2012; **206:** 58.e1–8.

58 Wright E, Audette MC, Ye XY, et al. Maternal vascular malperfusion and adverse perinatal outcomes in low-risk nulliparous women. *Obstet Gynecol* 2017; **130:** 1112–20.

59 Allen RE, Zamora J, Arroyo-Manzano D, et al. External validation of preexisting first trimester preeclampsia prediction models. *Eur J Obstet Gynecol Reprod Biol* 2017; **217:** 119–25.

60 Fraser A, Macdonald-Wallis C, Tilling K, et al. Cohort profile: the Avon Longitudinal Study of Parents and Children: ALSPAC mothers cohort. *Int J Epidemiol* 2013; **42:** 97–110.

61 Wright J, Small N, Raynor P, et al. Cohort profile: the Born in Bradford multi-ethnic family cohort study. *Int J Epidemiol* 2013; **42:** 978–91.

62 Cameroni I, Roncaglia N, Crippa I, et al. P32.05: Uterine artery Doppler in a risk population: what's its role in the prediction of severe pregnancy complications? *Ultrasound Obstet Gynecol* 2008; **32:** 421–22.

63 Carbillon L. The imbalance of circulating angiogenic/antiangiogenic factors is mild or absent in obese women destined to develop preeclampsia. *Hypertens Pregnancy* 2014; **33:** 524.

64 Carter J, Seed P, Tribe R, et al. Saliva progesterone for prediction of spontaneous preterm birth: the POPPY study. Pregnancy Outcome Poster Abstracts. *BJOG AN Int J Obstet Gynaecol* 2017; **124:** 122–54.

65 Figueiró-Filho EA, Oliveira VM, Coelho LR, Breda I. Marcadores séricos de trombofilias hereditárias e anticorpos antifosfolípides em gestantes com antecedentes de pré-eclâmpsia grave. *Rev Bras Ginecol Obstet* 2012; **34:** 40–46.

66 Giguère Y, Massé J, Thériault S, et al. Screening for pre-eclampsia early in pregnancy: performance of a multivariable model combining clinical characteristics and biochemical markers. *BJOG* 2015; **122:** 402–10.

67 Goetzinger KR, Singla A, Gerkowicz S, Dicke JM, Gray DL, Odibo AO. Predicting the risk of pre-eclampsia between 11 and 13 weeks' gestation by combining maternal characteristics and serum analytes, PAPP-A and free β-hCG. *Prenat Diagn* 2010; **30:** 1138–42.

68 Huang T, Hoffman B, Meschino W, Kingdom J, Okun N. Prediction of adverse pregnancy outcomes by combinations of first and second trimester biochemistry markers used in the routine prenatal screening of Down syndrome. *Prenat Diagn* 2010; **30:** 471–77.

69 Langenveld J, Buttinger A, van der Post J, Wolf H, Mol BW, Ganzevoort W. Recurrence risk and prediction of a delivery under 34 weeks of gestation after a history of a severe hypertensive disorder. *BJOG* 2011; **118:** 589–95.

70 Meertens LJE, Scheepers HC, De Vries RG, et al. External validation study of first trimester obstetric prediction models (Expect Study I): research protocol and population characteristics. *JMIR Res Protoc* 2017; **6:** e203.

71 North RA, McCowan LME, Dekker GA, et al. Clinical risk prediction for pre-eclampsia in nulliparous women: development of model in international prospective cohort. *BMJ* 2011; **342:** d1875.

72 Staff AC, Braekke K, Harsem NK, Lyberg T, Holthe MR. Circulating concentrations of sFlt1 (soluble fms-like tyrosine kinase 1) in fetal and maternal serum during pre-eclampsia. *Eur J Obstet Gynecol Reprod Biol* 2005; **122:** 33–39.

73 Stirrup OT, Khalil A, D'Antonio F, Thilaganathan B. Fetal growth reference ranges in twin pregnancy: analysis of the Southwest Thames Obstetric Research Collaborative (STORK) multiple pregnancy cohort. *Ultrasound Obstet Gynecol* 2015; **45:** 301–07.

74 Velauthar L, Plana MN, Kalidindi M, et al. First-trimester uterine artery Doppler and adverse pregnancy outcome: a meta-analysis involving 55,974 women. *Ultrasound Obstet Gynecol* 2014; **43:** 500–07.

75 Vollebregt KC, Gisolf J, Guelen I, Boer K, van Montfrans G, Wolf H. Limited accuracy of the hyperbaric index, ambulatory blood pressure and sphygmomanometry measurements in predicting gestational hypertension and preeclampsia. *J Hypertens* 2010; **28:** 127–34.

76 Caritis S, Sibai B, Hauth J, et al. Low-dose aspirin to prevent preeclampsia in women at high risk. *N Engl J Med* 1998; **338:** 701–05.

77 Coomarasamy A, Williams H, Truchanowicz E, et al. A randomized trial of progesterone in women with recurrent miscarriages. *N Engl J Med* 2015; **373:** 2141–48.

78 Dhillon-Smith RK, Middleton LJ, Sunner KK, et al. Levothyroxine in women with thyroid peroxidase antibodies before conception. *N Engl J Med* 2019; **380:** 1316–25.

79 Sibai BM, Caritis SN, Thom E, et al. Prevention of preeclampsia with low-dose aspirin in healthy, nulliparous pregnant women. *N Engl J Med* 1993; **329:** 1213–18.

80 Chappell LC, Seed PT, Briley AL, et al. Effect of antioxidants on the occurrence of pre-eclampsia in women at increased risk: a randomised trial. *Lancet* 1999; **354:** 810–16.

81 Coomarasamy A, Devall AJ, Cheed V, et al. A Randomized trial of progesterone in women with bleeding in early pregnancy. *N Engl J Med* 2019; **380:** 1815–24.

82 Makrides M, Gibson RA, McPhee AJ, et al. Effect of DHA supplementation during pregnancy on maternal depression and neurodevelopment of young children: a randomized controlled trial. *JAMA* 2010; **304:** 1675–83.

83 Poston L, Bell R, Croker H, et al. Effect of a behavioural intervention in obese pregnant women (the UPBEAT study): a multicentre, randomised controlled trial. *Lancet Diabetes Endocrinol* 2015; **3:** 767–77.

84 Rumbold AR, Crowther CA, Haslam RR, Dekker GA, Robinson JS. Vitamins C and E and the risks of preeclampsia and perinatal complications. *N Engl J Med* 2006; **354:** 1796–806.

85 Gordon NP, Lin TY, Rau J, Lo JC. Aggregation of Asian-American subgroups masks meaningful differences in health and health risks among Asian ethnicities: an electronic health record based cohort study. *BMC Public Health* 2019; **19:** 1551.

86 Lewis A. 'What group?' Studying Whites and Whiteness in the era of 'color-blindness'. *Sociol Theory* 2004; **22:** 623–46.

87 Romero R, Chaiworapongsa T, Erez O, et al. An imbalance between angiogenic and anti-angiogenic factors precedes fetal death in a subset of patients: results of a longitudinal study. *J Matern Fetal Neonatal Med* 2010; **23:** 1384–99.

88 Getahun D, Ananth CV, Kinzler WL. Risk factors for antepartum and intrapartum stillbirth: a population-based study. *Am J Obstet Gynecol* 2007; **196:** 499–507.

89 Flenady V, Koopmans L, Middleton P, et al. Major risk factors for stillbirth in high-income countries: a systematic review and meta-analysis. *Lancet* 2011; **377:** 1331–40.

90 Smith R, Mohapatra L, Hunter M, et al. A case for not adjusting birthweight customized standards for ethnicity: observations from a unique Australian cohort. *Am J Obstet Gynecol* 2019; **220:** 277.e1–10.

91 Hernández-Díaz S, Schisterman EF, Hernán MA. The birth weight "paradox" uncovered? *Am J Epidemiol* 2006; **164:** 1115–20.

92 Jackson M, Kiernan K, McLanahan S. Maternal education, changing family circumstances, and children's skill development in the United States and UK. *Ann Am Acad Pol Soc Sci* 2017; **674:** 59–84.

93 Mirowsky J, Ross CE. Education, socioeconomic status and health. In: Mirowsky J, Ross CE, eds. Education, social status and health, 1st edn. New York, NY: Routledge, 2017: 55.

94 Stewart AL, Dean ML, Gregorich SE, Brawarsky P, Haas JS. Race/ethnicity, socioeconomic status and the health of pregnant women. *J Health Psychol* 2007; **12:** 285–300.

95 Jackson MI. A life course perspective on child health, academic experiences and occupational skill qualifications in adulthood: evidence from a British cohort. *Soc Forces* 2010; **89:** 89–116.

96 Lorch SA, Enlow E. The role of social determinants in explaining racial/ethnic disparities in perinatal outcomes. *Pediatr Res* 2016; **79:** 141–47.

97 Hayward MD, Hummer RA, Sasson I. Trends and group differences in the association between educational attainment and U.S. adult mortality: implications for understanding education's causal influence. *Soc Sci Med* 2015; **127:** 8–18.

98 Jaspal R. Migration and identity processes among first generation British South Asians. *South Asian Diaspora* 2015; **7:** 79–96.

99 Haque HW. Factors influencing South Asian women's access to maternity related health services: a mixed methods study in an ethnically diverse urban setting in the UK. PhD thesis, University of East London. 2018. https://doi.org/10.15123/PUB.7801 (accessed June 30, 2022).

100 Public Health England. Public Health Outcomes Framework: Health Equity Report. Focus on ethnicity. July, 2017. https://assets.publishing.service.gov.uk/government/uploads/system/uploads/attachment_data/file/733093/PHOF_Health_Equity_Report.pdf (accessed June 30, 2022).

101 Elo IT, Vang Z, Culhane JF. Variation in birth outcomes by mother's country of birth among non-Hispanic black women in the United States. *Matern Child Health J* 2014; **18:** 2371–81.

102 UNICEF. Neonatal mortality. December, 2021. https://data.unicef.org/topic/child-survival/neonatal-mortality/ (accessed June 30, 2022).

103 Martin JA, Hamilton BE, Osterman MJ, Curtin SC, Matthews TJ. Births: final data for 2013. *Natl Vital Stat Rep* 2015; **64:** 1–65.

104 Gregory ECW, MacDorman MF, Martin JA. Trends in fetal and perinatal mortality in the United States, 2006–2012. *NCHS Data Brief* 2014; **2014:** 1–8.

105 Kingdon C, Roberts D, Turner MA, et al. Inequalities and stillbirth in the UK: a meta-narrative review. *BMJ Open* 2019; **9:** e029672.

106 McCall L. The complexity of intersectionality. *Signs* 2005; **30:** 1771–800.

107 Birthrights. Systemic racism, not broken bodies: an inquiry into racial injustice and human rights in UK maternity care. May, 2022. https://www.birthrights.org.uk/wp-content/uploads/2022/05/Birthrights-inquiry-systemic-racism-May-22-web-1.pdf (accessed June 30, 2022).

108 Waters A. Racism is "at the root" of inequities in UK maternity care, finds inquiry. *BMJ* 2022; **377:** o1300.

109 Pope R, Ganesh P, Miracle J, et al. Structural racism and risk of SARS-CoV-2 in pregnancy. *EClinicalMedicine* 2021; **37:** 100950.

110 Hamed S, Thapar-Björkert S, Bradby H, Ahlberg BM. Racism in European health care: structural violence and beyond. *Qual Health Res* 2020; **30:** 1662–73.

111 Aspen Institute. Racial equity: 11 terms you should know to better understand structural racism. July 11, 2016. https://www.aspeninstitute.org/blog-posts/structural-racism-definition/ (accessed June 30, 2022).

112 117th Congress, House Committee on Oversight and Reform. Birthing while Black: examining America's Black maternal health crisis. May 6, 2021. https://oversight.house.gov/legislation/hearings/birthing-while-black-examining-america-s-black-maternal-health-crisis (accessed June 30, 2022).

113 East CE, Biro MA, Fredericks S, Lau R. Support during pregnancy for women at increased risk of low birthweight babies. Cochrane Database Syst Rev 2019; published online April 1. https://doi.org/10.1002/14651858.CD000198.pub3.

114 Douglass C, Lokugamage A. Racial profiling for induction of labour: improving safety or perpetuating racism? *BMJ* 2021; **375:** n2562.

115 Royal College of Obstetricians and Gynaecologists. Race Equality Taskforce. 2020. https://www.rcog.org.uk/about-us/campaigning-and-opinions/race-equality-taskforce/ (accessed July 27, 2022).

116 Kapadia D, Zhang J, Salway S, et al. Ethnic inequalities in healthcare: a rapid evidence review. NHS Race and Health Observatory. February, 2022. https://www.nhsrho.org/wp-content/uploads/2022/02/RHO-Rapid-Review-Final-Report_v.7.pdf (accessed June 30, 2022)

117 NHS England. Improving equity and equality in maternity and neonatal care. Maternity Transformation Programme. 2022. https://www.england.nhs.uk/mat-transformation/improving-equity-and-equality-in-maternity-and-neonatal-care/ (accessed June 30, 2022).

118 NHS England. Core20PLUS5—an approach to reducing health inequalities. The Equality and Health Inequalities Hub. 2022. https://www.england.nhs.uk/about/equality/equality-hub/national-healthcare-inequalities-improvement-programme/core20plus5/ (accessed June 30, 2022).

119 Illinois Perinatal Quality Collaborative. Addressing maternal disparities and promoting birth equity. 2019. https://ilpqc.org/birthequity/ (accessed June 30, 2022).

120 New York State Department of Health. Taskforce on Maternal Mortality and Disparate Racial Outcomes. December, 2021. https://www.health.ny.gov/community/adults/women/task_force_maternal_mortality/ (accessed June 30, 2022).

121 Esegbona-Adeigbe S. The impact of a Eurocentric curriculum on racial disparities in maternal health. *Eur J Midwifery* 2021; **5:** 36.

122 NIHR. Promoting equality, diversity and inclusion in research. National Institute for Health and Care Research. 2022. https://www.nihr.ac.uk/about-us/our-key-priorities/equality-diversity-and-inclusion/ (accessed July 27, 2022).