

# MASTER

Optimal experiment design of a new minimal model of glucose kinetics during an oral glucose tolerance test

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# **Optimal Experiment Design** of a New Minimal model of Glucose **Kinetics** during an Oral Glucose Tolerance Test

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# Optimal Experiment Design of a New Minimal model of Glucose Kinetics during an Oral Glucose Tolerance Test

Incorporating effects of recruitment of muscle capillaries Glut-4 translocation and hepatic glucose production

S. Almagro Frutos, A.J.M. Wagenmakers, P.P.J. van den Bosch, N.A.W. van Riel

# Abstract

Investigating glucose kinetics under physiological conditions e.g. a meal or an oral glucose tolerance test (OGTT), is extremely valuable since diabetes mellitus is a pressing public health concern. In this study an expanded version of the glucose minimal model of Bergman has been developed and verified from a physiological and identification point of view using synthetic data produced from parameters found in literature.

In succession of the minimal model presented by [Bergman et al., Am. J. Phys. 236: E667-E677, 1979] which is widely used under non-physiological conditions, different minimal models are presented for use under physiological conditions during an OGTT e.g. [Natalucci et al., Int. Conf. IEEE EMB, Cancun, Mexico, 2003]. Natalucci's et al. minimal model is expanded to include *jointly* the effects of capillary recruitment, Glut-4 translocation and the production and inhibition of hepatic glucose production. The expanded minimal model incorporates a physiological model of glucose appearance from gut into blood plasma and an interstitial glucose compartment which can be measured using the microdialysis technique.

The new model has been tested in all its identifiability aspects: on a priori global identifiability by using an algorithm based on differential algebra [Audoly et al, *IEEE Trans Biomed.*, 48 1: 55-65, 2001] and a posteriori identifiability using model to model adjustment. Also, the sensitivity of the a priori identifiable parameters in relation with the estimated parameters has been examined showing large sensitivity for two of the six a priory parameters. Parameter estimation has been performed using Maximum likelihood estimation and the worst bias is calculated showing low bias for all estimated parameters. The oral glucose ingestion dosis has been optimised to reduce the bias for the estimated parameters and an algorithm has been designed to find the optimal sample schedule. Knowing the *joint* interaction between these mechanism in an novel, physiologically more detailed model and identifying the new oral glucose minimal model *in all its aspects and* performing optimal experiment design, is a great step forward to understand and quantify different potential causes of insulin resistance in type 2 diabetes.

Keywords: minimal compartmental model, glucose kinetics, OGTT, capillary recruitment, Glut-4 translocation, saturable receptor-mediated insulin transport, microdialysis, interstitial glucose, a priori global identifiability, a posteriori identifiability, maximum likelihood estimation, optimal experiment design

# Introduction

#### **I** Diabetes Mellitus

Diabetes mellitus is a metabolic disease characterized by varying or persistent elevated glucose levels in blood plasma. Diabetes mellitus is a pressing public health concern. In the Netherlands more than a half million people suffer from diabetes. Diabetes mellitus can be divided into two types: type I and type 2. Diabetes type I is caused by a lack of insulin

production and insulin secretion by the pancreas and type 2 diabetes is caused by decreased sensitivity of body tissues for insulin. This study focuses on type 2 diabetes.

The glucose regulatory system in humans consists of releasing insulin by the pancreatic  $\beta$  cells when the glucose concentration in blood rises above the normal range of 3.6 to 6.1 mM. Insulin stimulates glucose uptake by skeletal muscle (and adipose tissue) and suppress glucose production by the liver (hepatic glucose production). Insulin stimulates glucose uptake by skeletal muscle via two sequential mechanisms: recruitment of capillaries surrounding muscle fibers and glucose transporter-4 (Glut-4) translocation.

The second mechanism depends on the access of glucose and insulin to skeletal muscle. Glucose uptake by human skeletal muscle is regulated by insulin concentration in the interstitium [H. Herkner 2003]. Insulin is transported from blood plasma through the endothelium to reach the interstitium. After insulin has reached the interstitium, insulin is cleared due to internalization and degradation of insulin by muscle cells: Insulin binds to the insulin receptor in the interstitium which is the initiating step in insulin signalling pathways as discussed.

Two hypotheses have been proposed for the first mechanism: capillary recruitment and insulin-mediated increase of blood flow [Bergman 2003 and references in there]. In skeletal muscles two vascular flow routes exits: one route by nutritive capillaries that surround the skeletal muscles fibres responsible for the delivery of glucose and insulin and a second route by non-nutritive capillaries that has no contact with the skeletal muscles fibres. The first hypothesis suggests that insulin controls the blood flow to skeletal muscle at rest and during exercise by varying blood flow between the nutritive capillaries and the non-nutritive capillaries. It is believed that this process is very fast, 5 minutes after increase of insulin [Vincent 2004, Iwashita 2001, Porter 1997] and at insulin concentrations lower than required for the second hypothesis [Zhang 2004]. Changes in the distribution pattern can by caused by insulin resistance of the endothelial cells. Several studies [Clerk 2006, Gudbjörnsdóttir 2005, Gudbjörnsdóttir 2003] show an increase of capillary recruitment (increased permeability-surface area product (PS)) during enhanced physiological insulin concentration and a reduced capillary recruitment (lower PS) in type 2 diabetes humans than in healthy (obese) humans. The second hypothesis suggests that insulin acts as a vasolidator (alternating the blood flow at regular intervals through different capillary groups) controlling the access of glucose and insulin to skeletal muscle. Insulin binds to the insulin receptor on the luminal endothelial cells and this activates nitric oxide synthase (NOS). In response, endothelial cells release the vasodilator nitric oxide (NO) that prevents the smooth muscles cells of the precapillary arteriole to contract. This effect is more apparent at higher insulin concentration (pharmacologic concentrations) and after several hours. Impaired vasolidation can by due to insulin resistance. No increase in blood flow is seen during physiological insulin concentrations in obese healthy humans and type 2 diabetic humans [Bergman 2003] and reference in there, Gudbjörnsdóttir 2005, Gudbjörnsdóttir 2003]. However, [Clerk 2006] shows an increase during physiological insulin concentrations in lean healthy humans as well as in obese healthy humans. Insulin resistance can also have a more physical-chemical reason like a highly impermeable endothelium or reduced capillary density.

Several studies have adressed the mechanism of insulin transport from blood plasma through the endothelium to reach the interstitium by measuring glucose and insulin concentrations in blood plasma and interstitium. These studies result in conflicting conclusions either diffusion transport or a saturable receptor- mediated transport. These conflicting results have been analysed here (appendix A). It is concluded that a saturable receptor-mediated insulin transport is considered when insulin concentrations are in a physiological range (up to 600 pM) whereas a non-saturable insulin transport is considered at pharmacologic insulin concentrations (++ 600 pM).

The second mechanism by which insulin stimulates glucose clearance is by increased uptake by myocytes. The transport of glucose through the muscle cell membrane is facilitated by a special carrier mediated protein called Glut-4. In the early morning fasting state only a small portion of the total amount of Glut-4 is located in the outer muscle cell membrane (insulin independent). In response to the binding of insulin to the insulin receptor in the interstitium, triggering a cascade of protein phosphorylation reactions that ultimately activate the translocation of Glut-4 stored in microsomal membranes are transported to the outer muscle cell membrane. (insulin dependent). [Lauritzen 2006] shows a time delay of 10 minutes between insulin leaving the interstitium and the first translocations of Glut-4 after intravenous insulin infusion. Maximum translocation of Glut-4 was reached after 30 minutes.

Three impairments exist in the above described mechanisms for type 2 diabetic humans: Impairment of Glut-4 translocation, impairment of the recruitment of muscle capillaries, and impairment of the inhibition of hepatic glucose production [Zijlstra 2004]. The contribution of each of these impairments is not clear, since only an overall insulin sensitivity  $(S_i)$ , the efficacy of insulin to glucose uptake by body tissues  $[1/(pM \cdot min)]$ , is used in experiments to quantify the whole insulin resistance.

The goal of this study is to develop a method to understand and quantify different potential causes of insulin resistance in type 2 diabetes under normal physiological conditions. For this purpose a novel model is to be developed and to be verified. In general the model will incorporate the main glucose regulating mechanism to obtain descriptions of the variables involved and to enebale estimation of physiologically relavant parameters using measured data of glucose profiles, insulin profiles or others profiles of processes in a human.

The model should meet al the conditions for valid use in widespread clinical use. The oral glucose minimal model must describe the mechanism by which insulin regulates glucose uptake by muscle cells or adipose tissues as best as possible. Simplification of the model leads to undermodelling of the glucose regulatory system giving erroneous estimates. Overmodelling leads to complex model which are difficult to identify. The data from clinical measurements used for the novel model must be obtained during normal life conditions, the possibilities for invasive measurements in humans are limited and the costs of the must provide reliable unique estimates for the model parameters which describe the effects of Glut-4 translocation, the recruitment of muscle capillaries, and impairment of the inhibition of hepatic glucose production with high precision using the measurements techniques available, in other words the model needs to be identifiable for those parameters.

The first part of this report consists of modelling glucose kinetics in humans. First the mechanisms of glucose uptake by skeletal muscle and its possible impairments were discussed in this section. Further a closer analyse was made of the mechanism of transport of insulin from blood plasma through the endothelium to reach the interstitium list in appendix D. Secondly, several minimal models modelling the rate of appearance of glucose from the gut in the blood plasma are discussed. Thirdly, the new oral glucose minimal model is presented with is modifications compared to the classical minimal model of [Bergman 1979]. Finally the new oral glucose minimal model is represented in state space and the plasma

glucose, interstitial glucose, interstitial insulin and the rate of appearance of glucose are simulated using measured insulin profiles. A different approach could have been to use a model to simulate insulin (insulin profiles) that is secreted by pancreatic  $\beta$ -cells via the liver into the portal vein. A study discussing the model is list in appendix B.

The second part of this report consists of parameter estimation to identify the unknown parameters in the glucose metabolism model as optimal as possible. First, the glucose metabolism model is tested on a priori global identifiability and a posteriori identifiability. Secondly, the sensitivity of a priori parameters in relation with the estimated parameters is examined. Thirdly, the parameter estimation algorithm used is explained and also a technique is presented to calculate a lower bound of the precision that can be achieved for each parameter when estimating the parameters. Fourthly, an algorithm is designed to optimize the glucose input doses and an algorithm is designed to find the optimal sampling scheduling. Then the results are shown. Finally, a discussion is made and suggestions for future research are given.

#### 2 Minimal Models

 $S_i$  can be estimated in experiments with the Clamp technique, with the intravenous glucose tolerance test (IVGTT), with the meal oral glucose tolerance test (MGTT) or with the oral glucose tolerance test (OGTT). The Clamp technique consists of intravenous injecting a known insulin concentration together with an infusion of glucose to control the glucose concentration in blood plasma to a fixed value. The ratio of injected insulin and infused glucose provides a measure of insulin sensitivity. The IVGTT consists of injecting a known glucose concentration into the blood plasma. The glucose and insulin concentration in blood plasma are measured in time. The minimal model [R.N. Bergman 1979] is used to estimate parameters from which  $S_l$  is calculated. The CLAMP technique as well as the IVGTT use non-normal life conditions. The MGTT and OGTT are tests using normal life conditions. The MGTT uses an ingested meal containing glucose and the OGTT uses an ingested liquid containing glucose. The glucose and insulin profiles in blood plasma are measured. Also less surgical intervention is needed (only arterial and venous blood samples) than during a CLAMP or IVGTT. The minimal model of Bergman needs to be adapted to incorporate the rate of appearance of glucose from the gut in the blood plasma / systemic circulation of nutrient taken by mouth (Ra(t)).

In the past two mathematical models without physiological meaning have successfully been used to estimate Ra(t). The "integral equation" method is based on the assumption that Ra(t) is an anticipated version of blood plasma glucose concentrations [A. Caumo 2000]. The method calculates the areas under the measured glucose and insulin blood plasma concentrations. To properly calculate the areas, the glucose and insulin concentration have to return to basal levels. This may cause problems if the "integral equation" method is applied to humans with either impaired glucose tolerance or impaired insulin secretion, in which the return of glucose and insulin to the pretest levels requires a rather long time. Compared to the Bergman minimal model for an IVGTT, only two additional parameters need to be estimated. The "differential equation" method uses a piecewise-linear parametric model of Ra(t) [Dalla Man 5-2002]. An advantage of this method is that glucose and insulin concentrations do not have to return to basal value at the end of the test, hence, the experiment is dramatically shortened. Four additional parameters have to be estimated.

Two mathematical models with physiological meaning have been used successfully to estimate Ra(t). [D.R.L. Worthington 1997] shows a physiological model which was tested on a type I diabetic humans. A one compartmental method was proposed, having the following physiological parameters: fraction content of glucose in the food normalized to glucose distribution volume, time constant of gastric emptying and the transport delay between ingestion of glucose and appearance of glucose in blood plasma. Advantages of this method are that only three additional parameters have to be estimated and the parameters are physiological.

An interesting physiological model is studied by [J.D. Elashoff 1982]. This model is used in this study. The model uses the exponential curve which follows a set of physiological and mathematical characteristics. The most important mathematical characteristic is that at time zero the fraction of glucose ingested (*DS*) [g] is 1.0 and at infinite time the ingested fraction is 0.0. Further, the power exponential curve (1) also allows to have a slow initial emptying followed by an accelerated emptying ( $\beta$ >1). If  $\beta$ <1 the opposite occurs. This phenomenon is often seen in humans who have undergone an operation. If  $\beta$ =1 the exponential curve is obtained. Physiologically, the model uses the inverse of the time constant of gastric emptying (*k*) [1/min] and  $\beta$ .

$$f(t) = e^{(-k \cdot t)^{\beta}} \tag{1}$$

Data of olders studies obtained by measuring the contents of different ingested meals (e.g. glucose contents) and taken samples of the contents of stomach and using by different healthy humans as well as by healthy humans with operated stomach were fit to (I) using nonlinear least squares regression. This model showed high goodness of fit with the measured data.

Another study [J. Schirra 1996] performed an experiment with eight healthy non-obese males using this power exponential curve (1). The experiment used 50 or 100 grams glucose dissolved in 400 ml water and drank in 2 minutes showing k's and  $\beta$ 's with a very high precision.

Natalucci coupled the mono compartmental model [S. Natalucci 2003] using the power exponential curve (I) with the adapted minimal model [Bergman 1979] for an OGTT. Figure I shows this model were *Ggut* (t) [I/kg] is the mass of glucose in the gut and *Kabs* [I/min] is the rate constant of glucose absorption from the gut into blood plasma. The model was tested on nine non diabetic humans using a doses of 75 grams of glucose (*DS*) dissolved in water. However, the parameters of the power exponential curve were not estimated, but the values  $\beta$ =1.23 and *k*= 0.014 [I/min] were taken from [J. Schirra 1996]. The experiment showed *S*<sub>1</sub> and *Kabs* with low variance.



Figure 1: Mono compartmental model of the gut

#### 3 New Oral Glucose Minimal Model

Since the contribution of the different potential impairments in the glucose regulatory system in type 2 diabetes, as stated in the introduction, is not clear yet, we propose to split up  $S_I$  in the individual components  $S_I$ -cap,  $S_I$ -glut-4 and  $S_I$  –liver.  $S_I$ -cap  $[I/(pM \cdot min)]$  is the efficacy of insulin to stimulate insulin and glucose transport from blood plasma into interstitium by capillary recruitment.  $S_I$ -glut-4  $[I/(\mu M \cdot min)]$  is the efficacy of insulin to stimulate and adipose tissue by GLUT-4 translocation.  $S_I$ -liver  $[I/(pM \cdot min)]$  is the efficacy of insulin to inhibit hepatic production and stimulate glucose uptake by liver. The aim of this study was to modify the method developed by Bergman [R.N. Bergman 1979] to make this splitting possible.

[Zijlstra 2004] suggested a modification of the minimal model of Bergman. This modification consists of adding a (measured) interstitial glucose compartment to the classic minimal model, to model the control of glucose transport into the interstitium by insulin as discussed in the introduction Adding this compartment could make the identification of the rate constant of transport of plasma glucose into the interstitial compartment possible. This extra compartment is not new, a remote (non-accessible, only exchange with the glucose blood plasma compartment) glucose compartment was already used during an IVGTT in [Cobelli 1999], during an OGTT [Dalla Man 10-2002] and references in there to avoid overestimation of  $S_G$  (the insulin independent glucose uptake) and underestimation of  $S_I$  seen in one Compartmental Minimal Models (ICMM) like the classic Bergman model. This extra compartment was also used in [Caumo 1993] and references in there to derive a time profile of endogenous hepatic glucose production during a labeled IVGTT. However, this extra compartment was not measured in all earlier models.

Another difference in the new oral glucose minimal model with the classic minimal model of Bergman is that inhibition of hepatic production and stimulation of glucose uptake by liver is regulated by plasma insulin instead of interstitial insulin as access to hepatocytes is immediate, due to a highly fenestrated endothelium [Bergman 2003]. Also, plasma insulin regulates the transport of insulin and glucose from plasma to interstitium as discussed in the introduction.

In the classic minimal model of Bergman  $S_G$  consists of the net rate of glucose produced by the liver and insulin independent glucose uptake by tissues like the brain. Unfortunately, those effects cannot be estimated independently. By using ingested labelled glucose, a distinction is made between ingested (labelled) glucose and the net rate of (non-labelled) glucose produced by the liver. By measuring the concentration of labelled glucose in blood plasma  $S_I$ -liver can be obtained as e.g. performed in [Dalla Man 10-2002].

The suggestions made by Zijlstra [Zijlstra 2004] together with the new suggestions made in this section, the previously discussed physiologic description of glucose appearance from the gut into blood plasma and with the saturable receptor-mediated insulin transport hypothesis (appendix D) resulted in the new oral glucose minimal model 1 shown in figure 2.



Figure 2: New oral glucose minimal model for glucose metabolism, derived from the classic minimal model of Bergman including a interstitial glucose compartment G'(t), inhibition of hepatic production and stimulation of glucose uptake by liver regulated by plasma insulin I(t) instead of by interstitial insulin I'(t) and a physiologic description of glucose appearance from the gut into blood plasma G(t) and using a saturable receptor-mediated insulin transport from blood plasma to interstitium for insulin concentrations in physiological range

This new model consists of four compartments G(t) [mM], glucose plasma, G'(t) [mM], interstitial Glucose, I(t) [pM] insulin plasma and I'(t) [pM] interstitial insulin. Rge [g/min] is the rate of gastric emptying, DS [g] is the amount of glucose ingested, Ggut (t) [I/g] is the mass of glucose in the gut and Kabs [I/min] is the rate constant of glucose absorption from the gut into blood plasma.

Glucose G(t) is transported at a rate proportional to  $k_7$  [I/min] to the interstitium and cleared to muscle and adipose tissue at a rate proportional to  $k_1$  [I/min]. Glucose plasma G(t) is also cleared to tissue in a non-insulin dependent way at a rate proportional to  $p_6$  [I/min].

Insulin I(t) is transported at a rate proportional to  $k_2$  [1/min] to the interstitium. Insulin I(t) controls the inhibition and stimulation of hepatic glucose uptake by liver (k6 [1/( pM·min)]). Also, insulin controls the transport of plasma glucose (k8 [1/pM]) and plasma insulin (k9 [1/pM]) to the remote compartments. Interstitial insulin controls glucose uptake by muscles ( $k_4$  [1/( pM·min)]) and is degraded at a rate proportional to  $k_3$  [1/min].

Figure 3 shows the reparameterized model of figure 2 (model 1). In model 1, p2 equals k3 [1/min],  $S_1$ -cap equals  $k_2 \cdot k_9$  [1/(pM  $\cdot$  min)], p4 equals  $k_7 \cdot k_8$  [1/( pM  $\cdot$  min)], p5 equals k5 [1/min], p6 equals  $S_G$  [1/min],  $S_1$ -glut4 equals  $k_4 \cdot k_1$  [1/( pM  $\cdot$  min)] and  $S_1$ -liver equals  $k_6$  [1/( pM  $\cdot$  min)].



Figure 3: Reparameterized new oral glucose minimal model (model 1) for glucose metabolism

The reparameterized new oral glucose minimal model 1 of figure 2 can be modified further by including ingested labelled glucose GL(t) resulting in the new oral glucose minimal model 2 as shown in figure 4 (index L denotes the labelled glucose version of a compartment). Remark: Another method could be to enrich blood plasma with labelled glucose by injecting intravenous labelled glucose resulting in the new oral glucose minimal model 3. The new oral glucose minimal model 2 can easily be modified to describe this experimental setup.



Figure 4: New reparameterized oral glucose minimal model (model 2) for glucose metabolism including labelled ingested glucose

The new oral glucose minimal model describes *jointly* the effects of Glut-4 translocation, the recruitment of muscle capillaries, and impairment of the inhibition of hepatic glucose production using existing minimal models in literature and modifying them using recent insights in the mechanism by which insulin regulates glucose uptake by muscle cells or adipose tissue. Knowing the *joint* interaction between these mechanisms is a great step forward to understand and quantify different potential causes of insulin resistance in type 2 diabetes. Furthermore, analyzing and verifying the new oral glucose minimal model *in all its aspects* (a priori and a posterior identifiability, a priori parameter sensitivity) *and* performing optimal experiment design (optimal input and optimal sample schedule) is a great step forward in modelling glucose kinetics since to the authors knowledge it has never been applied all in one to a minimal model of glucose kinetics before.

# Methods

# 4.1 Measurements Techniques

The data will be obtained in the near future at the School of Sport and Exercise Sciences at the University of Birmingham. Sophisticated techniques are used to measure the concentrations of plasma glucose and insulin and interstitial glucose during an OGTT. Glucose and insulin in blood plasma are measured by taking samples by inserting an intravenous cannula (BD Venflon<sup>TM</sup>) into one of the antecubital veins. The subject's arm is placed in a so-called 'hot box' in which the high temperature will arterialise the venous blood. The blood samples will be centrifuged at 3000g at 4°C for 10 min and stored at -70°C for later analysis of glucose and insulin.

Plasma glucose concentrations will be determined by the Cobas Mira Plus automated analyser (ABX Diagnostics, UK) with a coefficient of variation (CV) of 2% if measured in 10-fold. Plasma Insulin concentrations will be determined by a commercially available kit (DRG instruments GmbH, Germany) with a CV of 10%. Further, the interstitial glucose concentrations are measured by the microdialysis technique using custom made CMA107 microdialyse probe with a length of 3 cm, a MW cut off of 20000 and a coefficient of variation (CV) of 4%. Details of this technique are discussed in the next section.

# 4.1.1 Microdialysis

The microdialysis technique [Lönnroth 1987, Lönnroth 1997] consists of the diffusion of a water-soluble substance over a dialysis membrane into a perfusion medium. It enables measurement of the interstitial fluid. The inlet of the catheter is perfused by a pump with a perfusion fluid (water containing isotonic saline) at a rate of  $0.667 \mu$ l/min. Exchange occurs between the perfusion fluid and the interstitial fluid. After equilibrium is reached, subsequent dialyses samples (20  $\mu$ l) are collected at the outlet of the catheter. The collected sample is a mean measurement over the time needed to reach equilibrium. The samples are analyzed for interstitial glucose. Due to the permeability of the dialysis membrane, maximal 85-90% of interstitial glucose is recovered in the dialysis. The microdialysis technique is not suitable for interstitial insulin measurements since very low recovery (+/- 3%) is achieved yet [Sjöstrand 2005, Sjöstrand 1999]. Another drawback of this method is that data samples

suffer from slow dynamics: interstitial mean glucose concentrations are only available after reaching equilibrium in about 30 minutes.

#### 4.2.1 Model Structure in State Space

The oral glucose minimal model I can be represented in state space. The state space model structure and its state space measurement model of a dynamic system is shown in equations (2, 3) and (4), respectively, where x is the *n*-state vector, u the multi input vector, f are linear or nonlinear functions which describe the structure of the system parameterized by vector  $\theta_i$ , g are nonlinear or linear functions which describe the output configuration parameterized by vector  $\theta_i$ ,  $\gamma_v [t_{iv}]$  are the measurements of the vth output sampled at  $N_v$  discrete times  $t_{iv}$ ,  $\gamma_{vm}$  (t,  $\theta_t$ ) is the modeled vth output at time t or sampled at  $N_v$  discrete times  $t_{iv}$ ,  $e_v$  is the measurement error, assumed to be additive zero mean white noise with known variance  $\sigma_v^2(t_i)$  and M the number of measured outputs.

State equation:

$$x(t,\theta_t) = f(x(t,\theta_t); u(t),\theta_t) \quad x(t_0,\theta_t) = x_0$$
<sup>(2)</sup>

Output equation:

$$y_{v}(t,\theta_{t}) = g_{v}(x(t,\theta_{t});\theta_{t})$$
(3)

Measurement model:

$$y_{\nu}[t_{i\nu}] = y_{\nu m}[t_{i\nu}, \theta] + e_{\nu}[t_{i\nu}] \quad i = 1, 2, 3, ..., N_{\nu}, \qquad \nu = 1, 2, 3, ..., M$$
(4)

This framework includes non-equidistant time samples, that not all outputs (blood plasma glucose and interstitial glucose) are sampled at the same time points and outputs measured with different accuracy. It should be noted that for the glucose minimal model two different types of input variables have to be discriminated. The orally ingested glucose can be manipulated by the experimentator and, therefore, is an independent input, which, in system theory, is commonly denoted by u(t). In our approach the measured insulin profile is also applied as an input to the minimal model. However, since this input is a variable which results from the response of the system to the independent input u(t), it is not independent.

The concept of 'dependent inputs' allows the opening of some of the feedback loops that connect the system of interest to the rest of the closed loop real system [van Riel 2006]. The model can focus on a smaller part of the system, as long as this subsystem is still integrated in its real system through the measured input signal. If the system contains a subsystem (state variables) which output can be measured, this subsystem (state variables) can be substituted by using the measured signal as input to the system (dependent input v(t)) as shown in (5). The number of states of the system is reduced. The drawback is a loss of predictive power, because the 'open-loop' model can only be used to simulate situations for which the dependent inputs have been measured in the real system. To solve the differential equations for the used integration step, the measured dependent input signal v(t) is linearly interpolated. If the dependent input(s) v(t) can be (accurately) measured in time, these signal can be used as forcing functions for these input(s).

$$x(t,\theta_t) = f(x(t,\theta_t); u(t), v(t), \theta_t) \quad x(t_0,\theta_t) = x_0$$
(5)

# 4.2.2 Simulation

The differential equations have been solved in Matlab by the differential equation solver Ode45. The relative error tolerances (RelTol) of the solutions were set four orders lower than the amplitudes of the outputs to  $1 \cdot 10^{-4}$  and the absolute error tolerances (AbsTol) were set two orders smaller than RelTol to  $1 \cdot 10^{-6}$ .

# 4.3 Identifiability

The parameters of the new oral glucose minimal model 1, 2 and 3 are tested on a priori global identifiability and a posteriori identifiability to analyse which parameters can be estimated using the measured techniques discussed in section 4.1. A parameter can be estimated if its variance has an acceptable low value.

A priori global identifiability consists of analyzing the structure of the model to verify which parameters can be estimated using the measured techniques discussed in section 4.1 and using noise-free data without having any a priori information about the parameters. In contrast to a priori global identifiability, a posteriori identifiability verifies which parameters can be estimated by analysing the effects of the dimensions of the parameters, the effects of different input profiles and the effects of different sample schedules.

# 4.3.1 A Priori Global Identifiability

A priori global identifiability under noise-free conditions has been examined by using an algorithm based on differential algebra [Audoly 2001]. Briefly, the algorithm consists of first deriving one equation which is algebraic observable, in other words may only contain the model input and output variables and their derivatives and not contain any states or its derivatives. The input-output differential equation representation of a dynamic model has this property. Hence, the differential equation has to be derived from the state space representation of the model. In [Audoly 2001] this is done by eliminating all states in each equation of the state space representation by substituting state equations or their derivatives in each other, the so called characteristic set of the model, resulting in one equation without states, the so-called input-output relation of the model

The differential equation of the model is identical to the input-output relation of the model. The differential equation is used to find the set of identifiable parameters by extracting the coefficients of the differential equation of the model. These coefficients are subtracted by the same coefficients which are substituted by a random numerical point in the parameter space for the unknown parameters. An example is shown below:

Example A:

Differential equation:

$$c1 \cdot c2 \cdot x(t) + c3/c2 \cdot x(t) = 0$$
 (a1)

Coefficients of the differential equation:

(a2)

A random numerical point in the parameter space for the unknown parameters:

$$c1 = 12$$
 (a3)  
 $c2 = 4$   
 $c3 = 36$ 

Coefficients of the differential equation subtracted by the same coefficients substituted by random numerical point in the parameter space for the unknown parameter:

$$c1 \cdot c2 - 48 = 0$$
 (a4)

$$c_3/c_2 - 9 = 0$$
 (a5)

The obtained set of coefficients (example 1: (a4, a5)) of the differential equation, which are called exhaustive summary in [Audoly 2001], are solved by the Buchberger algorithm [Buchberger 2001] obtaining the Gröbner basis as shown in (a6). If this basis has a finite number of solutions, all parameter can be estimated properly. Else, non-identifiable parameters have to be estimated jointly, to obtain a finite number of solutions of the Gröbner basis. In example 1 C1 and C2/C3 are identifiable.

$$9c2-c3$$
  
-9+c3/c2  
 $c1-48/c2$  (a6)

The algorithm described in [Audoly 2001] assumes generic initials conditions for the state space model. Generic initials conditions are initial equations which are not restricted to a particular value. If non-generic initials conditions are used, parameters could not be observable in the differential equation leading to parameters which cannot be tested on identifiability using this algoirtum. Hence, if non-generic initials conditions are used a manual check needs to be done to verify if unknown parameters are not observable when deriving the input-output relation of the model. If parameters vanish, the set of state equations of the state space representation needs to be enlarged by adding the polynomials evaluated at these initial conditions. In other words, the initial condition is filled in the state space representation of the model. The input-output relation must be recalculated.

[Saccomani 2002, Saccomani 2001] presented a modified algorithm based on [Audoly 2001] for dealing with those non-generic initial conditions. The algorithm is based on the lack of reachability / accessibility from the initial state. If the system is accessible from all initial states, the characteristic set of the model is valid. Else the characteristic set needs to be enlarged adding the polynomials evaluated at the initial conditions which lacks of accessibility. The accessibility property of a model is checked by using the accessibility Lie Algebra (6) where  $\partial/\partial x$  denotes the Jacobian matrix, A are the polynomials of the state space model which are multiplied by the states x and B are the polynomials of the state space model which are multiplied by the input u as defined in (7). A model is accessible from its initial conditions if the so-called accessibility rank condition holds for each initial condition [Sontag 1998]. The accessibility rank condition is satisfied when the linear span of all vectors functions that can be obtained starting with the A 's and taking any numbers of Lie brackets of them as shown in (8) has the same dimension as the state space model.

$$[A,B] = \left(\frac{\partial}{\partial x}B\right)A - \left(\frac{\partial}{\partial x}A\right)B \tag{6}$$

$$x = A \cdot x + B \cdot u \tag{7}$$

 $[A,B],[A,[A,B]],[A,[A,[A,B]],[A,[A,[A,[A,[A,B]]],...=[AB],[A^{2}B],[A^{3}B],[A^{4}B],...$ (8)

# 4.3.2 A posteriori Identifiability

The model to model adjustment [Damen 2005] approach has been used to analyse the a posteriori identifiability using different input profiles under noise and noise-free conditions and under. The effect using when using different samples schedules can also be studied. Model to model adjustment consists of analyzing the error cost function (9) for combinations of parameters  $\tilde{\theta}$  with different values in the parameter space where  $\gamma_{\nu}(t_{i\nu})$  are the measurements of the  $\nu$ th output sampled at  $N_{\nu}$  discrete times  $t_{i\nu}$ ,  $\gamma_{\nu m}(t_{i\nu},\theta_t)$  are the modeled  $\nu$ th output sampled at  $N_{\nu}$  discrete times  $t_{i\nu}$  using the parameters  $\theta_i$ ,  $z_{\nu}$  is the  $\nu$ th output weighting factor,  $\theta_t$  are the true parameters,  $N_{\nu}$  the number of samples taken for each output  $\nu$  and M the number of outputs.

$$J(\tilde{\theta}) = \frac{1}{N_{v}} \sum_{i=1}^{N} \sum_{v=1}^{M} z_{v} \left( y_{v}(t_{iv}, \theta_{i}) - y_{vm}(t_{iv}, \tilde{\theta}) \right)^{2}$$
(9)

The 2 dimensional contour plot, plots the cost function (the model error) for two parameters for a range of values. Other parameters in the parameter space are kept fixed at the true parameter  $\theta_t$ . If the combination of parameter is identifiable, the contour plot shows concentric ellipses centered at the supposed true values of the two parameters (global minima). The more narrow and steeper the concentric ellipsis are, how more accurate the estimation of both parameters will be.

#### 4.3.3 A Priori Parameters Sensitivity

Parameters which are not identifiable by the used model setup in combination with the used experimental setup have been set a priori. Unfortunately, parameters can vary among population. The parameter sensitivity shows if a non-identifiable parameter can be set a priori without losing significant precision of the parameters when the non-identifiable parameter varies significantly among population. Therefore the parameters sensitivities have been calculated by varying the a priori parameters one by one over a fixed range with respect to the true known parameter, setting the value of the other a priori parameters fixed, estimating the corresponding unknown identiafiable parameters and finally calculating the corresponding parameter covariance matrix  $\Delta(\theta)$  which will be explained in section 4.4. If the identifiable parameters are unbiased, the determinants of the obtained covariance matrices are an indication for the parameter sensitivity for variation of the a priori parameter. Else, if estimates were not unbiased, the worstcase bias technique, as stated in the last paragraph of section 4.4, has been used.

# **4.4 Parameter Estimation**

Several algorithms exist to estimate parameters of a model, such as: Weight Least Squares estimation (WLS), Maximum Likelihood estimation (ML), Generalized Least Squares estimation and Bayesian estimation (MAP). In this study Maximum Likelihood estimation has been used since the noise introduced by the measurements techniques used (section 4.1) can be assumed to be white, Gaussian, zero mean and with variance  $\sigma_v^2$ .

The ML criterion assumes that the measured vth output  $y_{\nu}$  sampled at  $N_{\nu}$  discrete times  $t_{i\nu}$  consists of the modelled vth output  $y_{\nu m}$  depending on true parameters  $\theta_t$  and noise *e* as shown in (10). which is different from (4) since the parameter vector has been split in a priori unidentifiable parameters which were kept fixed  $\theta_{fix}$  and the identifiable parameters which have been estimated  $\theta_{est}$ .

$$y_{\nu}[t_{i\nu}] = y_{\nu m}[t_{i\nu}, \theta_{fix}, \theta_{est}] + e_{\nu}[t_{i\nu}] \quad i = 1, 2, 3, ..., N_{\nu}, \quad \nu = 1, 2, 3, ..., M$$
(10)

The probability density function of Gaussian noise is shown in (10).

$$p_{v}(e_{v}(t)) = \frac{1}{\sqrt{2\pi}\sigma_{v}}e^{-\frac{e_{v}(t)^{2}}{2\sigma_{v}^{2}}}$$
(11)

Combining equations (10) and (11) and assuming that the consecutive noise samples are independent, results in the likelihood function for the residue of the vth output with variance  $\sigma_{\nu}^{2}$  (12).

$$L(y_{v} \mid \boldsymbol{\theta}_{est}) = \left(\frac{1}{\sqrt{2\pi\sigma_{v}}}\right)^{N_{v}} e^{-\sum_{i=1}^{N_{v}} \frac{(y_{v}(i_{iv}) - y_{vm}(i_{iv}, \boldsymbol{\theta}_{fis}, \boldsymbol{\theta}_{eu}))^{2}}{2\sigma_{v}^{2}}}$$
(12)

The noise is assumed to be white, Gaussian, zero mean and with variance  $\sigma_{\nu}^{2}$ . This assumption has been checked, as will be discussed in this section, after having estimated the parameters.

The likelihood function for the residues of all outputs *M* is shown in (13).

$$L(y \mid \theta_{est}) = \prod_{\nu=1}^{M} \left( \frac{1}{\sqrt{2\pi}\sigma_{\nu}} \right)^{N_{\nu}} e^{-\sum_{i=1}^{N_{\nu}} \sum_{\nu=1}^{M} \frac{\left(y_{\nu}(t_{i\nu}) - y_{\nu m}(t_{i\nu}, \theta_{fix}, \theta_{est})\right)^{2}}{2\sigma_{\nu}^{2}}$$
(13)

Taking the log of (13) results in the cost function for the ML (14) [Damen 2005 p. 47-62].

$$J(\theta_{est}) = \sum_{i=1}^{N_v} \sum_{\nu=1}^{M} \frac{\left(y_{\nu}(t_{i\nu}) - y_{\nu m}(t_{i\nu}, \theta_{fix}, \theta_{est})\right)^2}{2\sigma_{\nu}^2}$$
(14)

The model parameters  $\theta_{est}$  are estimated by minimizing (14) as shown in (18).

$$\theta_{est} = \underset{\theta_{est}}{\arg\min} J(\theta_{est})$$
(15)

Function minimization of (15) has been accomplished in Matlab using the non linear least square estimation algorithm lsqnonlin which uses the Medium-Scale Gauss-Newton Algorithm with Line search cubicpoly. The minimization of (15) stopped when the variation in cost function, between two subsequent iterations was less than TolFun or the absolute parameters value changed less than TolX. TolFun was set much lower than the amplitudes of the weighted outputs  $z_v y_v$  to  $1 \cdot 10^{-4}$  and TolX, a factor 1000 lower than the lowest true value of the estimated parameter set to  $1 \cdot 10^{-10}$ . The minimization algorithm also stops after a number of iterations is reached. The maximal number of iterations was set to 800.

After having estimated the parameters, the residues of the criterium  $J_{\nu}(\theta_{est})$  must be analyzed to ensure that the assumption made earlier that the noise is Gaussian, zero mean with variance  $\sigma_{\nu}^{2}$  is true. For white noise it holds that consecutive noise samples are independent. Its mean is zero and its autocorrelation function  $\phi_{\nu}(t_{shift})$  (16) takes the form of a pulse at zero time shift and is zero elsewhere (Dirac pulse).

$$\phi_{v}(t_{shift}) = \frac{1}{N_{v}} \sum_{i=1}^{N_{v}} J_{v}(t_{iv}, \theta_{est}) \cdot J_{v}(t_{iv} + t_{shift}, \theta_{est})$$
(16)

The lower bound of the parameter covariance matrix  $\Delta(\theta_{est})$  of the estimated model parameters  $\theta_{est}$  is obtained (using the Cramer-Rao theorem) from the inverse of the Fisher information matrix  $F(\theta_{est})$  [Walter 1997 p. 245-250] as shown in (17). If MLE qualifies as an unbiased and minimum variance estimator, the lower bound in (17) becomes an equality.

An estimator is unbiased if (18) holds where E denotes the expectation. In addition, if the variance  $\sigma_v^2$  is set to the inverse of de data covariance of the disturbance and N<sub>v</sub>>>n<sub>0</sub> (number of parameters) holds, MLE qualifies as an asymptotically unbiased and minimum variance estimator [Damen 2005], respectively.

$$\lim_{N \to \infty} \Delta(\theta_{est}) = F^{-1}(\theta_{est}) = \sum_{\nu=1}^{M} \left( \frac{E}{y_{\nu m} | \theta_{est}} \left\{ \frac{\partial^2 \left( -\ln\left(L\left(y_{\nu m} + \theta_{fix}, \theta_{est}\right)\right)\right)}{\partial \theta_{est} \partial \theta_{est}} \right\} \right)^{-1}$$

$$\approx \sum_{\nu=1}^{M} \left( \frac{1}{N_{\nu}} \sum_{i=1}^{N_{\nu}} \frac{1}{2\sigma_{\nu}^{2}} \frac{\partial y_{\nu m}(t_{i\nu}, \theta_{fix}, \theta_{est})^{2}}{\partial \theta_{est}} \frac{\partial y_{\nu m}(t_{i\nu}, \theta_{fix}, \theta_{est})^{2}}{\partial \theta_{est}} \frac{\partial y_{\nu m}(t_{i\nu}, \theta_{fix}, \theta_{est})^{2}}{\partial \theta_{est}} \right)^{-1}$$

$$(17)$$

$$E(\theta_{est} - \theta_t) = 0 \text{ for } N_{\nu} >> n_{\theta}$$

(18)

with *E* the expected value and  $n_{\theta}$  the number of estimated parameters. The variance  $\sigma_v^2$  in (17) has been determined a posteriori. After having estimated the parameters, information is available concerning the variance  $\sigma_v^2$ . It is set to the squared 2-norm of the residues of  $J(\theta_{ry})$  (14).

If the parameters are biased  $(E(\theta_{est} - \theta_t) \neq 0)$  the parameter covariance matrix cannot be calculated using (17). In this case, the relative worst case bias is used to provide a measure of the (in)accuracy of the parameter estimations. The relative worst case bias as shown in (19) has been calculated for each identifiable parameter  $\theta_{est}$  in the parameter space by varying the initial values of each identifiable parameter one  $\theta_{est}$  by one between -50 to +50% of  $\theta_t$  in steps of 10%, setting all other identifiable parameter fixed and estimating its value using the maximum likelihood estimation technique and adding for each estimation procedure q a different sequence of noise to the synthetic data with a 1/SNR of 0%, 5% or 10%. The relative worst case bias is then the maximal difference found, among all Q estimation procedures, between estimate  $\theta_{est,q}$  and true value  $\theta_{q,t}$  as shown in (20).

$$Bias_{rel}(\%)_{est,q} = \frac{\left|\theta_{est,q} - \theta_{t}\right|}{\theta_{t}} \cdot 100\% \qquad q = 1, 2, 3 ..., Q$$
(19)

$$Bias_{rel}(\mathscr{H})_{est} = M_{\theta} x \left( Bias_{rel}(\mathscr{H})_{est,q} \right)$$
(20)

# 4.5 Optimal Experiment Design

Designing an optimal experiment for parameter identification consists primary of selecting the inputs and outputs to make all parameters of interest in the model identifiable (section 4.3). Secondary, it consist of choosing a suitable parameter estimation algorithm, optimizing the shape of the input signal, choosing the optimal sampling schedule, choosing the type of sensors etc. to get the maximal information from the obtained data.

Since the choices of inputs and outputs were restricted due to physical accessibility, optimal experiment design has been focused on optimizing the glucose input doses and choosing the optimal sampling schedule discussed in section 4.5.1 and 4.5.2 respectively.

# 4.5.1 Optimal Input Shaping

An optimal input can be determined by the D-optimality criterion [Walter 1997 p. 287-291] as applied in [Cobelli 1988] if the identifiable parameters are unbiased. D-optimality criterion consists of minimizing the determinant of the parameter covariance matrix, which is equivalent to maximizing the determinant of the Fisher information matrix (17) (unbiased and minimum variance estimator) for the input signal u(t) using a conventional sample schedule that is *intuitively* used by experts. The optimal input criterion for each vth output is shown in (21).

$$\begin{split} &\underset{u(t)}{\text{MIN}} \{ \det[\Delta(\theta_{est}, u(t))] \} = \underset{u(t)}{\text{MAX}} \{ \det\left[F(\theta_{est}, u(t))\right] \} = \underset{u(t)}{\text{MAX}} \{ \det\left[-\sum_{\nu=1}^{M} \underset{y_{\nu m}}{E} \left\{ \frac{\partial^{2} \ln(L(y_{\nu m} \mid \theta_{fix}, \theta_{est}, u(t))))}{\partial \theta_{est} \partial \theta_{est}^{\mathsf{T}}} \right\} \right\} \end{split}$$

$$\approx \underset{u(t)}{\text{MAX}} \{ \det\left[\left(\frac{1}{N_{\nu}} \sum_{i=1}^{N_{\nu}} \frac{1}{2\sigma_{\nu}^{2}} \frac{\partial(y_{\nu m}(t_{i\nu}, \theta_{fix}, \theta_{est}, u(t)))^{2}}{\partial \theta_{est}} \frac{\partial(y_{\nu m}(t_{i\nu}, \theta_{fix}, \theta_{est}, u(t)))^{2}}{\partial \theta_{est}} \frac{\partial(y_{\nu m}(t_{i\nu}, \theta_{fix}, \theta_{est}, u(t)))^{2}}{\partial \theta_{est}} \right\} \} \end{split}$$

$$(21)$$

If the identifiable parameters are biased, (17) cannot be used to estimate the variance of the parameters. In this case, the relative worst case bias technique has been used as measure of parameter accuracy, as discussed in the last paragraph of section 4.4.

# 4.5.2 Optimal Sample Schedule

The parameter covariance matrix  $\Delta(\theta_{est})$  also depends on the used sample schedule according to the inverse of the Fisher information matrix (17).

The optimal sample schedule for each output is determined by maximizing the determinant of the Fisher matrix (17) (D-optimality) (unbiased and minimum variance estimator) for each sample taken  $(t_{i\nu})$  if the identifiable parameters are unbiased. The optimal sample schedule criterion for each vth output is shown in (22).

$$\begin{split} &MIN\left\{\det[\Delta(\theta_{est}, t_{iv})]\right\} \\ &= MAX\left\{\det[F(\theta_{est}, t_{iv})]\right\} = MAX_{t_{iv}}\left\{\det\left(-\sum_{v=1}^{M} \sum_{y_{vm} \in P, t_i}^{E} \left\{\frac{\partial^2 \ln(L(y_{vm} \mid \theta_{fix}, \theta_{est}, t_{iv})))}{\partial \theta_{est}}\right\}\right\}\right\} \\ &\approx MAX\left\{\det\left(\sum_{v=1}^{M} \left(\frac{1}{N_v}\sum_{i=1}^{N_v} \frac{1}{2\sigma_v^2} \frac{\partial(y_m(\theta_{fix}, t_{iv}, \theta_{est}, t_{iv}))^2}{\partial \theta_{est}} \frac{\partial(y_m(\theta_{fix}, t_{iv}, \theta_{est}, t_{iv}))^2}{\partial \theta_{est}}\right)\right\}\right\} \end{split}$$
(22)

If the parameters are biased the Fisher information matrix cannot be used. In this case the relative worst case bias, as stated in the last paragraph of section 4.4, has been used to provide the variance of the parameter estimations.

Several algorithms are based on starting with a conventional sample schedule (the sample schedule which is used normally) where each individual sample is optimized by comparing it with its neighbouring samples in the conventional sample schedule. If by taking a neighbouring sample, the determinant of the Fisher matrix is increased, this new sample is exchanged with the old one [Mori 1979], [Landaw 1982], [Xianjin 2000].

The algorithm developed in this study starts at the first sample of the conventional sample schedule searching for the first sample within a fixed horizon of samples that increases the precision. This sample is exchanged with the old one. This process is repeated until the maximal duration of the experiment is achieved. If after finishing, the optimal sample schedule contains too many samples than ethical/financial possible, the fixed horizon of samples is increased. Changing the horizon of samples is stopped when the optimal sample schedule contains the desired number of samples.

The search horizon was taken arbitrary to reduce the number of samples taken: twice the time needed to measure interstitial glucose. The optimal samples have been selected in the range (domain) of the conventional samples

The outputs do not have to be sampled at the same time points. The samples are separated by at least the time needed to take a measurement of the output. Measurement of G(t) takes much less than one minute and measurement of G'(t) takes thirty minutes.

# Results

#### 5.1.1 Model Structure in State Space

The new oral glucose minimal model I, valid for insulin concentrations in a physioligal range (up to 600 pM), is a multi input multi output (MIMO) nonlinear dynamic system consisting of a four state-vector (26-29) having one independent input, one dependent input and two outputs. The independent input is DS [g], the dependent input I(t) [pM] and the outputs (30-31) are  $G(t,\theta_i)$  [mM] and  $\overline{G'(t,\theta_i)}$  [mM] a mean over each 30 minutes as stated in section 4.1.2. The dependent input could become an independent input if simulated using the  $\beta$ -cell insulin secretion models which are list in Appendix B. The outputs are sampled at discrete time  $t_i$  with  $N_G$ ,  $N_{G'}$  samples, respectively, as shown in (32, 33) were  $G_m[t_i, \theta_{fix}, \theta_{est}]$  is the modelled quantized glucose ouput and  $\overline{G'_m[t_j, \theta_{fix}, \theta_{est}]}$  is the modelled quantized mean interstial glucose output over each 30 minutes. Due to measurement restrictions, the minimum time intervals of samples of  $G(t,\theta_i)$  and  $\overline{G'(t,\theta_i)}$  were taken as 1 minute and 30 minutes, respectively.  $e_v$  is the measurement error, assumed to be additive zero mean white noise with known variance  $\sigma_v^2$  for output v. The remaining parameters are  $G_b$  [mM] which is the basal glucoce concentration,  $I_b$  [pM] the basal insulin concentration,  $\sigma_1^2$  the variance of the measurement noise and  $\sigma_2^2$  the variance of the measurement noise.

$$\mathbf{\dot{G}}(t) = -[S_{I-liver} + p4] \cdot [G(t) - G_B] \cdot [I(t) - I_b] + p5 + S_G \cdot G(t) \cdot \frac{Kabs \cdot G_{gut}(t)}{V}$$

$$G(0) = G_b$$

$$(26)$$

$$\dot{I}'(t) = -p2 \cdot I'(t) + S_{1-cap} \cdot [I(t) - I_b]^2$$

$$I'(0) = 0$$
(27)

$$\dot{G}'(t) = -S_{I-g/ut-4} \cdot I'(t) \cdot G'(t) + p4 \cdot [G(t) - G_B] \cdot [I(t) - I_b] \qquad G'(0) = G'_b \qquad (28)$$

$$y_1(t, P_t) = G(t, \theta_t)$$
(30)

$$y_2(t, P_t) = G'(t, \theta_t)$$
(31)

$$y_{1}[t_{i}] = G_{m}[t_{i}, \theta_{fix}, \theta_{est}] + e_{v}[t_{i}] \quad i = 1, 2, 3, ..., N_{G}.$$
(32)

$$y_{2}[t_{j}] = G'_{m}[t_{j}, \theta_{fix}, \theta_{est}] + e_{v}[t_{j}] \quad j = 1, 2, 3, ..., N_{G}.$$
(33)

# 5.1.2 Simulation

Synthetic datasets are produced by simulating the model states space representation in Matlab using parameters found in literature for healthy non-obese humans and measured insulin input profile from [Wagenmakers 2006]. Sixth datasets were produced. Details of the datasets and plots used are shown in appendix C. Plots of the simulated glucose and insulin plasma and interstitial glucose data are also found in appendix C.

The parameters k, Kabs,  $\beta$ , and DS related by the gut were estimated in [Natalucci 2003] and [Schirra 1996] as stated in section 3, the parameters  $p_5$  and  $S_1$ -liver related by the liver were estimated by the two compartment minimal model (using a non-accessible interstitial glucose compartment, but having similar behaviour as a accessible interstitial glucose compartment) [Dalla Man 10-2002] and [Caumo 1993], the parameters p2, p6 and V were estimated by the (improved) hot IVGTT two compartment minimal model [Toffolo 2003], the parameter k7 was estimated by [Regittnig 2003] using high frequently sampled ISF (openflow microperfusion) which uses a mathematical model [Steil 1996] to describe transendothelial transport of glucose form blood plasma to remote interstitium and finally k2 was estimated by experimental data of [Sjöstrand 2005] using low frequently sampled ISF (mircrodialysis) and a mathematical model [Steil 1996] to describe transendothelial transport of insulin form blood plasma to remote interstitium. Mean values and variances are shown in table 1. No a priori information of the parameters k8, k9 and  $S_1$ -glut4 were found in literature. Parameters k8 and k9 have been, to the authors knowledge, never been estimated. Since the product of  $k_7/k_2$ ,  $k_8/k_9$  and I(t) must be in the range of the transport from blood plasma to interstitium  $k_7/k_8$ , these parameters were first settled roughly, proportional to the inverse of the concentration of insulin to maintain the same range of transport. Secondly, parameters k8 and k9 were adjusted by trial and error producing a possible dataset of glucose kinetis.  $S_{I}$ glut4 was totally unknown and was set by the trial and error method just discussed.

Kabs	Beta	k	V	k9	p2	S <sub>I</sub> -cap
2.89 · 10 <sup>-2</sup>	1.23	0.014	1.386	$\pm 2.0 \cdot 10^3 \cdot$	0.113	k2 · k0
$\pm 0.92 \cdot 10^{-2}$	± 0.09	± 0.001	$\pm 3.7 \cdot 10^{-2}$	(I/0.2I) <sup>3 III</sup>	± 0.023	<b>112</b> 119
S <sub>G</sub>	k2	k8	S <sub>1</sub> -glut4	k7	S <sub>I</sub> -liver	P4
0.018	0.0395	+ 1 0 10-2 [	+1.0.10-2	(0.063±7·10 <sup>-3</sup> )	1.57 • 10-6	k=*k8
± 0.006	± 0.008	± 3.0.10	±1.0.10	·(1/0.21) <sup>III</sup>	± 1.08 · 10 <sup>-6</sup> 1	k/ k0

Table 1: Overview of the parameter value used in the new oral glucose minimal model 1 to simulate the sixth datasets (appendix C)

<sup>&</sup>lt;sup>1</sup> Conversion factor between [ml/min  $1/\mu$ U] to [1/min 1/pM]: 1/7.0625, <sup>2</sup> Conversion factor between [dl/mg] to [1/min 1/mM]: 180.16 1/mM, <sup>3</sup> Ratio between distribution volume of blood plasma and interstitial fluid.

# 5.2 Identifiability

In the new oral glucose minimal model I two outputs are measured namely, measurement of plasma glucose G(t) and interstitial glucose G'(t). A restriction exits for measurements of interstitial glucose G'(t), as only a mean measurement is available over each 30 minutes  $(\overline{G'(t)})$  as stated in section 4.1.2. This restriction is taken into consideration by calculating the mean of  $\overline{G_m(t)}'$  over 30 minutes using the parameters  $\theta_i$  in the oral minimal glucose model. The mean is calculated by summing up the values of  $G_m'(t)$  with intervals set to one-fifth of the smallest experimental time constant used in the model and dividing it by the number of intervals.

#### 5.2.1 A Priori Global Identifiability

A priori global identifiability is done for the new oral glucose minimal models 1, 2 and 3. First the manual check is performed to verify if unknown parameters are not observable when dealing with non-generic initial conditions. Fortunately all parameters are observable. Table 2 shows the identifiable parameters.

	Identifiable parameter sets
model 1	Ib, Gb, kabs, v, p4, and Sicap or Siglut4
model 2	Ib, Gb, kabs, v, p4, SiLiver or p5 and Sicap or Siglut4
model 3	Ib, Gb, kabs, v, p4, SiLiver or p5, Beta or k and Sicap or Siglut4

Table 2: Overview of the a priori global identifiability parameters for models 1, 2 and 3

A remark can be made about the identifiability of  $S_1$ -cap.  $S_1$ -cap can be identified separately if the interstitial insulin G'(t) profile is measured and  $p_2$  is known a priori. Unfortunately, measurements of interstitial insulin G'(t) are not available. However, the interstitial insulin profile can be estimated after estimating  $S_1$ -glut4. In this case the variance of  $S_1$ -cap. depends on a summation of two parts. One part is its own variance when estimating it from the measured plasma insulin profile and estimated interstitial insulin. The other part is non linear and depends on the variance of  $S_1$ -glut4.

#### 5.2.2 A Posteriori Identifiability

The cost function of the new oral glucose minimal model I is evaluated for the plasma glucose output signal and the interstitial glucose output signal. The cost function is shown in (34).

$$J(\theta) = \frac{1}{N_G} \cdot \sum_{i=1}^{N_G} \left( G(t_i) - G_m(t_i, \theta) \right)^2 + \frac{1}{N_G} \cdot \sum_{i=1}^{N_G} \left( \overline{G'(t_i)} - \overline{G_m'(t_i, \theta)} \right)^2$$
(34)

For each combination of the two dimensions in the parameter space, the contour plot of the error function is plot for model 1 (DS=75 gram), 2 (DS=72 gram, labelled ingested glucose 3

gram) and 3 (DS=75 gram, enrichment labelled glucose 3%) with variations of the parameter  $\theta_t$  of +/- 50 % using steps of 5%. The contour lines for  $J(\theta) <= 10$  are plot at 0.01 distance from each other. For  $J(\theta) > 10 <= 200$ , the contour lines are plot at 5.00 distance from each other. The contour plots also include estimation results indicated by red circles. (The true value is indicated with a red plus-sign.) The a posteriori identifiability is examined for the first and the second data set described in section 5.1.2 / appendix C with added white noise having a 1/SNR of 0 % and 5%. Plots, not all, are shown in Appendix D.

Appendix D shows for model I using dataset I, that the contourplots for all possible combinations of parameters kabs,  $p_4$ ,  $S_1$ -glut4 and Gb are narrow and steep centered at the supposed true values of the two parameters (global minima). In contrast, parameter Ib shows for the same combinations of parameters in almost all case heavily stretched ellipses indicating poor-identiafiable parameters. Parameter  $\nu$  shows stretched ellipses for the same combinations with kabs and  $S_1$ -glut4.  $S_1$ -liver and all other parameters not included in appendix D.I or D.2 show heavily stretched ellipses for combinations with kabs,  $p_4$ , and  $S_1$ -glut4. The same holds when a different insulin profile (dataset 2) is used.

Model 2 is examined by the model 2 model adjustment to know whether ingesting labelled glucose using the same experimental setup improves identifiability of  $S_I$ -liver. However,  $S_I$ -liver show heavily stretched ellipses for combinations with kabs,  $p_4$ ,  $S_I$ -glut4 and Gb. Although model 3 shows less stretched ellipses than model 2 for combinations with kabs,  $p_4$ ,  $S_I$ -glut4 and Gb,  $S_I$ -liver remains a poor-identiafiable parameter.

Although a priori global identifiability shows identifiability of *beta* and *k* in all possible combinations with *kabs*,  $p_4$ ,  $S_1$ -glut\_4 and Gb in model 3, stretched ellipses are shown for all possible combination in appendix D.3.

In all cases that no noise is added, parameters are estimated exactly (unbiased parameters). By adding a I/SNR of 5% white noise to the ouitput signal, bias was introduced in the estimation of parameters. However, parameter estimations are all close to the exact value (not shown in appendix D).

# 5.2.3 A Priori Parameters Sensitivity

The a priori parameters sensitivity is calculated using the first and the second data set described in section 5.1.2/ appendix C with added white noise having a I/SNR of 0%, 5% and 10%, respectively. The a priori parameters are varied one by one by  $\pm 5$ , 10, 20 and 30%. The initial values of the identifiable parameters used in the estimation algorithm are set to those which showed largest bias in the relative worst case bias test in the case of a I/SNR of 0% white noise (section 5.3). The results shown in appendix E, shows the best possible variance for each identifiable parameter when the a priori parameters varies form the true value used.

Appendix E shows that for all estimated parameters, *Ib*, *Gb*, *beta* and *k* are highly sensitivity to variations of the true value among population. Parameter  $p_2$  only shows highly sensitivity for parameter  $S_1$ -glut4. The remaining parameters  $p_5$ ,  $S_G$  and  $S_1$ -liver are insensible. An exception occurs for  $S_1$ -liver which shows a quite high sensitivity when white noise is added with a I/SNR.of 10%.

## 5.3 Parameter Estimation

The cost function for ML is shown in (35) where  $\sigma_G^2$ ,  $\sigma_{G'}^2$  are the variance of measurement errors in plasma glucose G(t) and interstitial glucose G'(t).

$$J(\boldsymbol{\theta}, i) = \frac{\sum_{i} \left( G(t_i) - G_{vm}(t_i, \boldsymbol{\theta}_{fix}, \boldsymbol{\theta}_{est}) \right)^2}{2\sigma_G^2} + \frac{\sum_{i} \left( \overline{G'(t_i)} - \overline{G'_m(t_i, \boldsymbol{\theta}_{fix}, \boldsymbol{\theta}_{est})} \right)^2}{2\sigma_{G'}^2}$$
(35)

A priori information is available about the quality of the measurements of glucose and insulin plasma and interstitial glucose specified by the manufacturer of the measurement equipment or the lab where the samples are analyzed as stated in section 4.1.

The relative worst case bias is first calculated using the first and the second data set described in section 5.1.2 / appendix C with added white noise having a I/SNR of 0%, (2.5%), 5% and 10%, respectively. The initial values of the identifiable parameters used in the estimation algorithm are varied between +/- 50 % using steps of 10%. The results shown in table 3, shows the relative worst case bias for each identifiable parameter.

I/SNR Insulin profile I						Insulin	profile 2	
(%)	kabs	v	P4	SiGlut4	kabs	v	P4	SiGlut4
0	3.31	3.76	2.51	2.49	2.29	3.36	2.53	2.63
5	1.18	2.64	2.46	3.66	0.412	5.95	5.28	6.53
IO	2.40	6.26	6.19	4.43	0.34	13.0	10.5	_ 12.4

i/SNR	Insulin profile 1	Insulin profile 2
(%)	SiCap	SiCap
0	3.00 · 10-3	0.201
2.5	0.351	1.21
5	5.09	4.09
10	31.5	39.5

Table 3: Overview of the worst case bias of all identifiable parameters using synthetic data for two different insulin profiles (data sets 1 and 2, respectively, appendix b) with added white noise having a 1/SNR of 0, (2.5), 5 and 10%

Table 3 shows that all estimated parameters suffer from bias in all cases. All estimated parameters shows a low variance in all cases. (The initial values of the identifiable parameters in the case of a 1/SNR of 0% white noise which showed largest relative worst case bias are for the first dataset  $1.4 \cdot kabs$ ,  $1.4 \cdot v$ ,  $1.4 \cdot p$   $S_1$ -cap,  $1.5 \cdot S_1$ -glut4 and  $1.3 \cdot S_1$ -cap, for the second dataset  $0.5 \cdot kabs$ ,  $0.5 \cdot v$ ,  $0.5 \cdot p_4$ ,  $0.5 \cdot S_1$ -glut4 and  $0.6 \cdot S_1$ -cap.)

Secondly, the relative worst case bias is calculated using the third, fourth, fifth and sixth data set described in section 5.1.2/ appendix C with added white noise having a 1/SNR of 0%, 5% and 10%, respectively. The results are shown for datasets third and fourth in table 3 and for datasets fifth and sixth in table 4.

1/SNR	SNR Insulin profile 1					Insulin	profile 2	
(%)	kabs	v	P4	SiGlut4	kabs	V	P4	SiGlut4
0	18.2	17.1	16.6	4.89	16.5	20.0	19.7	12.8
5	17.4	28.1	16.6	24.2	24.7	23.1	25.4	5.20
10	20.6	38.8	36.4	48.9	60.1	18.4	23.8	10.5

1/SNR				Insulin	profile 2			
(%)	kabs	v	P4	SiGlut <sub>4</sub>	kabs	v	P4	SiGlut4
0	28.3	27.5	29.9	13.0	8.52	8.09	30.8	4.68
5	1.18	9.4I	69.2	18.0	10.9	9.39	49.9	20.8
IO	20.3	9.69	52.2	35.5	11.1	11.2	52.4	20.9

Table 4: Overview of the worst case bias of all identifiable parameters using synthetic data for two different insulin profiles and two different sample schedules (data sets 3, 4, 5 and 6, respectively, appendix b) with added white noise having a I/SNR of 0, 5 and 10%

Table 4 shows for all estimated parameters a higher variance for each estimated parameter than by using the first and second dataset, but still acceptable. However, table 4 shows a lower variance for all estimated parameters in the case of a 1/SNR of 0% and 5% than by using the third and fourth dataset but still higher than by using the first and second dataset.

# 5.4 Optimal Experiment Design

# 5.4.1 Optimal Input Shaping

The optimal dosis was calculated by varying the input dose DS in the range of I gram to 100 gram in steps of I gram using the first and the second data set described in 5.1.2/ appendix C with added white noise having a I/SNR of 0%, 5% and 10%, respectively. The initial values of the identifiable parameters used in the estimation algorithm were set to those which showed largest bias in the worst biased test in the case of a I/SNR of 0% white noise (section 5.3).

If the relative worst bias of each estimated parameter was lower than 10% and did not increase for an increase in glucose dosis, this *minimum* glucose input was said to be optimal. The results are shown in table 5.

1/SNR	Insulin profile 1	Insulin profile 2
(%)	Amount of glucose [g]	Amount of glucose [g]
0	I	T
5	33	33
10	59	78

Table 5: Optimal dosis using synthetic data for two different insulin profiles (data sets 1, 2, respectively, appendix B) with added white noise having a 1/SNR of 0, 5 and 10%

# **6** Discussion and Suggestion

The goal of this study was to investigate and validate a new oral glucose minimal model describing jointly the effects of Glut-4 translocation, the recruitment of muscle capillaries and the effects of impairments of the inhibition of hepatic glucose production using existing minimal models in literature and modifying them using recent insights in the mechanism by which insulin regulates glucose uptake by muscle cells or adipose tissues. The new oral glucose minimal model was to be identified in all its aspects (a priori and a posteriori identifiability, a priori parameter sensitivity and to perform optimal experiment design).

Three oral glucose minimal models have been developed. The first model was developed to estimate the rate constant of glucose absorption from the gut into blood plasma (*kabs*), to estimate the efficacy of insulin to stimulate its own transport from blood plasma into interstitium by capillary recruitment ( $S_1$ -cap), to estimate the efficacy of insulin to stimulate glucose transport from blood plasma into interstitium ( $p_4$ ), to estimate the efficacy of interstitial insulin to stimulate glucose clearance into muscle by Glut-4 translocation ( $S_1$ -glut\_4) and finally, to estimate the plasma glucose distribution volume per unit of body weight ( $\nu$ ). The second model was developed to estimate the previous parameters, but also to estimate the efficacy of insulin to inhibit hepatic production and stimulate glucose uptake by liver ( $S_1$ -liver) by including ingested labelled glucose. The third model was an alternative for the second model since model 3 enriches blood plasma with labelled glucose through injecting labelled glucose instead of ingesting labelled glucose.

After developing the three oral glucose minimal models, a priori global identifiability analysis proved that with the models structure used, these estimations were possible. However, more information was acquired when choosing model 3, since model 3 could also estimate *beta*, the variable of the power exponential curve denoting the velocity of emptying of the gut. This was obvious since extra data was available to estimate  $\nu$  by the infusion of labelled glucose.

A posteriori identifiability analysis showed, for all possible combinations of parameters *kabs*,  $p_4$ ,  $S_1$ -glut\_4 and Gb, contour plots with narrow ellips that were steep centered at the supposed true values of the parameters (global minima) using two different insulin profiles. Since Gb could be measured, this parameter was set a priori in order to not influence negatively the parameter estimation of the other unknown parameters. Parameter  $\nu$  showed stretched ellipses when estimated together with parameters *kabs* or  $S_1$ -glut\_4 indicating low accuracy and less reliability as was seen when the relative worst bias was calculated. A problem occurs if parameter  $S_1$ -liver was estimated since all three models showed heavily stretched ellipses for combinations of *kabs*,  $p_4$ , and  $S_1$ -glut\_4. Heavily stretched ellipses indicate non-identifiable parameters. Parameter  $S_1$ -liver in (26).

Further, the a priori parameters sensitivity was calculated using two different insulin profiles. All estimated parameters, *Ib*, *Gb*, *beta* and *k* showed highly sensitivity to variations of the true value among population. Since *Ib* and *Gb* could be measured, this was not a problem. On the other hand, high a priori parameter sensitivity of parameters *beta* and *k* was a real problem. A priori global identifiability showed identifiability of either *beta* or *k*, however, analyzing the a posteriori identifiability rejected this possibility. Good estimates of *kabs* and *k* are available in literature [Schirra 1996], however, study [J.D. Elashoff 1982] shows that humans who have undergone an operation in the stomach can be modelled by changing parameter *beta*. Hence, variations of *beta* among populations exist. *Beta* or *k* could be estimated by measuring the content of glucose in the gut or by using high(er) concentrations of labelled glucose. Parameter *p2* showed only highly sensitivity for parameter *S*<sub>1</sub>-glut4. Estimates of *p2* are

available in literature e.g. [Toffolo 2003]. The bias of  $S_I$ -glut4 could be lowered by using the open-flow microperfusion technique [Bodenlenz 2005] instead of the microdialysis technique for measurements of interstitial glucose  $\overline{G_m(t)}$ '. The open-flow microperfusion technique makes faster sampling of interstitial glucose  $\overline{G_m(t)}$ ' possible than the microdialysis technique. The remaining parameters  $p_5$ ,  $S_G$  and  $S_I$ -liver were insensible for variation of the a priori identifiable parameters. An exception occured for  $S_I$ -liver that showed a quite high sensitivity when noise was added with a I/SNR of 10% to the ouput signals. However, a maximal I/SNR of 5 % was present in the measured outputs using the microdialysis technique.

A remark could be made about the identifiability of  $S_1$ -cap.  $S_1$ -cap could be identified separately if the interstitial insulin G'(t) profile was measured and  $p_2$  was known a priori. Unfortunately, measurements of interstitial insulin G'(t) were not available. However, the interstitial insulin profile could be estimated after estimating  $S_1$ -glut4. In this case the variance of  $S_1$ -cap depends on a summation of two parts. One part is its own variance after estimating it from both the measured plasma insulin profile and the estimated interstitial insulin. The other part is non-linear and depends on the variance of  $S_1$ -glut4.

The studied identifiability agreed with the found relative worst case bias. A relative worst case bias of maximal 6.5 % was achieved for all estimated parameters using a 1/SNR of 5% added white noise. If a 1/SNR of 10% white noise was added, a maximal relative worst case bias of 13% was achieved. If no noise was added, a maximal relative worst case bias of 3.76% was achieved indicating bad working of the minimalization algoritm used to minimize (15). Important to note is that measurements of interstitial glucose G'(t) had been undermodelled. The interstitial glucose G'(t) samples were modelled by calculating the mean of  $\overline{G_m(t)}$  over 30 minutes by summing up the values of  $G_m'(t)$  with intervals set to one-fifth of the smallest experimental time constant used in the model and dividing it by the number of intervals. In fact, subsequent samples of interstitial glucose  $\overline{G_m(t)}$  are influenced by each other since the samples are collected at the outlet of the catheter without stopping the inlet of new interstitial glucose, causing mixing of subsequent interstitial glucose samples. This effect was not taken into account since the behaviour of mixing is still unknown.

The optimal dosis was calculated by varying the input doses *DS* in the range of I gram to 100 gram in steps of I gram using two different insulin profiles and using sequences of added white noise with a I/SNR of 0%, 5% and 10%. A complication occured when *DS* was varied. If *DS* vary, this also implies a different insulin profile. However, the same insulin profile was used during all variations of *DS* since an insulin profile of glucose ingestion other than 75 or 100 gram was not available. It is not common to have such an OGTT. This complication set doubts to the obtained results. This *minimum* glucose input was said to be optimal if the relative worst bias of each estimated parameter was lower than 10% and did not increase for an increase in glucose doses. For each 1/SNR a different optimum glucose input was obtained. For a 1/SNR of 0% of added white noise, all glucose doses were optimal. For a 1/SNR of 5% of added white noise, 78 gram was obtained. It could be concluded that the doses of glucose should be as high as possible.

Due to limited time issues, the optimal sample schedule algorithm was not applied. Instead of this, a different test was applied: The relative worst case bias was calculated for two extra

different sample schedules (appendix C). The first extra sample schedule (the conventional sample schedule) is used as standard in many studies [Sjostrand 2005], [Wagenmakers 2006]. For the second extra sample schedule, a new sample schedule was developed which takes samples of glucose and insulin plasma during events when most changes take places in the insulin profile. This is often, during the first 60 minutes after glucose ingestion. This second extra sample schedule showed a lower variance for all estimated parameters in the case of a 1/SNR of 0% and 5% of added white noise than using the first extra sample schedule. This is obvious, because the second extra sample schedule has samples of glucose and insulin plasma during events when the most changes takes places in the insulin profile. Hence, it is very interested to find the optimal sample schedule to obtain the lowest possible relative worst bias for the estimated parameters. In both cases, the relative worst case bias was higher than the best relative worst case bias calculated in section 5.3 because fewer samples are taken in both extra sample schedules. It could be discussed if an optimal samples schedule could be found, since great variety exits among measured glucose profiles.

A remark could be made about the optimal sample schedule algorithm developed here to calculate the optimal sample schedule. The optimal sample schedule could use the Fisher information matrix as indication how optimal a samples schedule is. At the start of the algorithm a sample schedule with many samples is available which showed, as in section 5.3, low relative worst case bias. However, as the algorithm progress, the *current* optimum sample schedule has fewer samples than the *initial* sample schedule, introducing more bias in the estimation of parameters. Hence, the conditions of Fisher information matrix are not satisfied, and the relative worst case bias technique must be used instead. A drawback of the relative worst case bias technique is the processing time which is extremely large. But if the search of the optimum sample schedule progress well, the sample schedule should always show low bias and the Fisher information matrix could be used. A dilemma that must be investigated, before using the optimal sample schedule algorithm.

A different remark can be made about the choice of first calculating the optimal input doses before searching for the optimal sample schedule or quite the reverse. Since the optimal samples is selected in the range (domain) of the conventional samples, it is obvious to calculate first the optimal input doses for the conventional samples and then to reduce this set of samples obtaining the optimal sample schedule.

#### 7 Conclusion

The goal of this study was to investigate possible impairments of insulin resistance in type 2 diabetes. Model 1 could be used to investigate whether there are impairments in Glut-4 translocation and impairments in the recruitment of muscle capillaries during physiological insulin concentrations that. The model did not take into account insulin production by the pancreas, instead of this the insulin concentration profile in blood plasma wasmeasured which insulin was produced by the pancreas. Model 1 could not discriminate impairment of hepatic glucose production.

The following parameters can be estimated using model 2 and measuring glucose blood plasma, insulin blood plasma, basal glucose level, basal insulin level and interstitial glucose using the microdialysis technique: The rate constant of glucose appearance from the gut into blood plasma, the efficacy of insulin to stimulate its own transport from blood plasma into interstitium by capillary recruitment, the efficacy of insulin to stimulate glucose transport from blood plasma into interstitium, the efficacy of interstitial insulin to stimulate glucose clearance into muscle by Glut-4 translocation and the plasma glucose distribution volume per unit of body weight. This was the first time that these parameters were estimated jointly to understand and quantify different potential causes (mechanism of Glut-4 translocation and recruitment of muscle capillaries) of insulin resistance in type 2 diabetes during an OGTT using a physiological model of glucose appearance in blood plasma. This model refines earlier developed models.

Using a minimum glucose doses of 78 gram during an OGTT and having a I/SNR of 10%, parameters were estimated with a maximum variance of  $\pm 13\%$  if each minute a sample of insulin and glucose blood plasma and each 30 minutes a sample of interstitial glucose were taken during the first 240 minutes after glucose ingestion. If the I/SNR is 5%, parameters were estimated with a maximum variance of  $\pm 6.5\%$ . An exception holds for the parameter which denotes the efficacy of insulin to stimulate its own transport from blood plasma into interstitium by capillary recruitment which can be slight higher. This is due to the fact that this parameter also depends in a still unknown non-linear way on the precision of the efficacy of interstitial insulin to stimulate glucose clearance into muscle by Glut-4 translocation. If fewer samples are taken, it is advised to take samples when most dynamics takes places in glucose kinetics or to use the optimal sample schedule algorithm developed in this study.

All a priori identifiable parameters used in model I except *beta* and *k* are insensitive for the estimated parameters when having a I/SNR of 0%, 5% and 10% of white noise and using model I and the experimental setup described in this study. Varying the true value of all the a priori identifiable parameters (except *beta* and *k*) found in literature with  $\pm$  30% in steps of 5% resulted in a maximum variance of  $\pm$ 12% in the estimation of the identifiable parameters. On the other hand *beta* and *k* showed a maximum variance of respectively 241% and 195%. Although, having this high sensitivity for the a priori identifiable parameters *beta* and *k*, the description of glucose appearance with model I is better than earlier developed models found in literature. The earlier developed models were based on interpolating glucose appearance in blood plasma instead of using physiological parameters to model glucose appearance in blood plasma.

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# Modelling the insulin transport from blood plasma through the endothelium to reach the interstitium.

Several studies have studied the mechanism of insulin transport from blood plasma through the endothelium to reach the interstitium concluding conflicting results: a saturable receptor-mediated or diffusion transport. A saturable receptor-mediated is possible since [R.S. Bar 1978] found receptors on cultured human endothelial cells obtained from human umbilical veins and [M.L. Peacock 1982] found receptors on cultured and isolated bovine endothelial cells obtained from pulmonary and systemic vessel.

Some in vivo studies have measured the insulin concentrations in the interstitium with the microdialysis technique [Lönnroth 1987, Lönnroth 1997] in human skeletal muscle of conscious healthy lean humans with no medication in (semi) supine position during an OGTT [H. Herkner 2003, M. Sjöstrand 2005] and during a CLAMP [M. Sjöstrand 1999, P.E. Jansson 1993]. Other studies have measured insulin concentration using the peripheral lymphatic cannulation technique [Castillo 1991] in hindlimb lymph of conscious healthy humans with a wide range of obesity and with no medication in (semi) supine position during a CLAMP [C. Castillo 1994] and with anesthetized healthy dogs during a CLAMP [M. Hamilton-Wessler 2002, G. M. Steil 1996]. Another study has measured insulin concentration in thoracic duct lymph with conscious healthy dogs during a CLAMP [Y.J. Yang 1989]. (In all lymph based studies, lymph fluid concentration is increased by massaging or saline injection.) No difference exists between the two different measurements sites, since [Y.J. Yang 1989, C. Castillo 1994] showed strong correlation and direct proportionality between insulin concentration in the hindlimb lymph near the human skeletal muscle and glucose uptake by human skeletal muscle. Only a delay time exits between insulin in the interstitium and insulin in hindlimb lymph due to the transport time needed between both places. From this, it can be assumed that insulin concentration in interstitium and hindlimb lymph also shows strong correlation and direct proportionality.

Saturable receptor-mediated transport of insulin can be shown by an increase of insulin gradient (or decrease of insulin ratio) between blood plasma and interstitium/lymph at higher insulin concentration. After all, an increase of insulin gradient (or decrease of insulin ratio) implies a decrease in insulin transport indicating a saturable receptor-mediated transport of insulin. But since changes in clearance of insulin by muscle cells affect the gradient, it is critical to assess this parameter if an increase in the gradient is to be interpreted as a decrease in transport. However, studies [Y.J. Yang 1989, C. Castillo 1994, H. Herkner 2003] show a strong correlation between glucose uptake and insulin concentrations in interstitium/lymph concluding that no saturation occurs in clearance of insulin by muscle cells. Hence, analyzing this gradient is significant for the mechanism of insulin transport through the endothelium.

After analyzing the discussed gradient in experimental data of several in vivo studies an interesting conclusion is drawn. The studies using insulin concentrations at physiological range (up to 600 pM in blood plasma) [Y.J. Yang 1989, H. Herkner 2003, C. Castillo 1994, M. Sjöstrand 2005] show a decreased gradient indicating saturable receptor-mediated transport. [P.E. Jansson 1993] shows at insulin concentrations of 3500 pM a deceased gradient indicating saturable receptor-mediated transport. The studies [G. M. Steil 1996, M. Hamilton-Wessler 2002] using insulin concentrations at pharmacologic range (more than 5000 pM in blood plasma) show an increased gradient indicating diffusion transport. As a unification of both conclusions, a saturable receptor-mediated insulin transport is considered under physiological conditions whereas a non-saturable insulin transport is considered at pharmacologic insulin concentrations.

In vitro study (cultured bovine aortic endothelial cells with the use of dual chamber, not exposed to a flow) [L. King 1985] at physiological range concludes a receptor-mediated transport of insulin through the endothelium. Another study in vitro (endothelial cells from rat hearts) [F. Brunner 1998] at physiological range concludes a diffusion process. However, [M.A. Vincent 2005] shows a different behavior of the endothelium of rats compared with the endothelium of humans. The insulin

concentration ratio between interstitium and blood plasma is around 80% in contrast to humans 40-50% [M.A. Vincent 2005, M. Sjöstrand 1999] or dogs 35-50% [M.A. Vincent 2005]. Hence, the endothelium in rats is not a good representation of the endothelium in humans. In vitro study (cultured endothelial cells in a hollow fiber apparatus and continuously exposed to a flow) [F. Salvetti 2002] at physiological range shows a decreased gradient at higher insulin concentrations. Since these studies are all in vitro, the conclusions drawn from them are taken less serious than in vivo studies. However, the conclusions are in agreement with in vivo conclusions.

Hence, the classic minimal model of Bergman [Bergman 1979] is only valid using insulin concentrations at physiological range. This is always the case during an OGTT.

Another interesting case to study, are possible delays. A significant delay about 20 minutes in the activation phase of transporting insulin from blood plasma to interstitium is seen in vivo studies [M. Sjöstrand 2005, M. Sjöstrand 2002, C. Castillo 1994, Y.J. Yang 1989, E. Jansson 1993] and in vitro study [F. Salvetti 2002]. The cause of this delay is not known yet. No delay is seen in the deactivation phase of transporting insulin from blood plasma to interstitium [Y.J. Yang 1989]. A small delay is seen between glucose uptake and lymph insulin for insulin binding to its cognate cell receptor [Y.J. Yang 1989].

# Modelling insulin that is secreted by pancreatic $\beta$ -cells via the liver into the portal vein.

The ability to control glucose production and utilization depends on insulin that is secreted by pancreatic  $\beta$ -cells via the liver, where insulin is decreased (extracted) by approximately 50%, into the portal vein. When the concentration of glucose in plasma increases rapidly, insulin is secreted by pancreatic  $\beta$ -cells in two phases: the first and second phase. The first 10 minutes (first phase) is a fast rise of insulin secretion. The next 20 minutes (second phase) is a gradually rise of insulin secretion which is directly related to the degree and duration of the stimulus [Caumo 2004]. Between the 70s and 80's several  $\beta$ -cell insulin secretion models were developed based on measurement of plasma insulin concentrations or measurement of c-peptide. These models can be added to the new oral glucose minimal model discussed in section 3 to estimate the plasma insulin profile instead of measuring the plasma insulin profile. Adding one of these models into the new oral glucose minimal model discussed on measurements of plasma insulin concentrations are given and second point  $\beta$  to based on c-peptide measurements of plasma insulin concentrations are given.

Toffolo proposed eight models tested on conscious healthy lean dogs by an IVGT [Toffolo 1980] using the minimal model [Bergman 1979]. The models assume that clearance of insulin is of the first order, the initial insulin peak represents a bolus loaded into insulin plasma after the glucose injection and the rate of the secondary rise in insulin is determined by the concentration of glucose in plasma above a specific threshold value. This model has the disadvantage that parameter estimation has to be done in two steps. First the insulin concentration is used as input data to estimate the parameters of the minimal model [Bergman 1979], secondly the recorded glucose data is used as input for the  $\beta$ -cell insulin secretion model. However, by the single-step fitting process [Zheng 2005] this handicap is resolved.

This proposed minimal model by Toffolo in 1980 is tested on healthy lean and obese humans by an IVGT [Bergman 1981] showing strong correlation with [Toffolo 1980]. This study also shows that lean lower tolerance humans is related to 77% diminished second phase responsivity ( $S_1$  remained unchanged) and that obese lower tolerance humans 60% diminished  $S_1$  (second phase remained unchanged).

Zheng introduced a modification of Toffolo's model [Zheng 2005] by the assumption that the insulin decay rate is not always a first-order process showing improved fitting (97.7%) than Toffolo's model in 1980. He also added a mathematical function for describing the endogenous insulin infusion rate. The model is shown in (b1) where  $\gamma$  [IpM·1/(mM·min)] is the rate of pancreatic release of insulin after the glucose injection, h [mM] is the pancreatic threshold value, t [min] is the time, n [(IpM)<sup>1-m</sup>·1/min] is the plasma insulin decay rate constant, u(t) [mU/mg·min] is the endogenous insulin infusion rate, m stands for the *m*-th order processes and k [pM] is the initial plasma insulin concentration after the glucose bolus injection. Due to the assumption that that the insulin decay rate is not always a first-order process equation is modified by (b2).

$$I(t) = \begin{cases} \gamma \cdot [G(t) - h] \cdot t - n \cdot [I(t) - I_b]^m + u(t)/V & \text{if } G(t) > h \\ -n \cdot [I(t) - I_b]^m + u(t)/V & \text{if } G(t) \le h \end{cases}$$
  $I(0) = k + I_b (bt)$ 

$$I'(t) = -p2 \cdot I'(t) + p2 \cdot S_{I} \cdot [I(t) - I_{b}]^{m} \qquad I(0) = 0$$
(b2)
Steil introduced a new model [Steil 2003] which assumes a proportional component that reacts to the difference between plasma and basal glucose (*P*), an integral component that adjusts basal delivery in proportion to hypo/hyperglycemia (*I*) and a derivative component that responds to the rate of change in glucose plasma (*D*). The model is test during a CLAMP on healthy humans showing good fitness. Furthermore, this model is especially designed for closed loop systems linking a glucose sensor with an insulin pomp (an automated portable continuous glucose monitoring system) that includes stability and improved frequency characteristics/response. The model is shown in (b3, b4) where *SR(t)* [pM] is the  $\beta$ -cell insulin secretion rate, *n* [I/min] is the reciprocal of the insulin plasma decay rate constant, K<sub>p</sub> [pmol·1/min·1/M] is the rate of  $\beta$ -cell insulin secretion, T<sub>I</sub> [min] is the rate of proportional to integral release and T<sub>p</sub> [Min] is the ratio of derivative to proportional release.

$$\begin{split} \dot{I}(t) &= -n \cdot I(t) + \frac{SR(t)}{V} \\ SR(t) &= P(t) + I(t) + D(t) \\ P(t) &= K_p \cdot (G(t) - G_b) \\ \dot{I}(t) &= K_p \cdot (G(t) - G_b) / T_i \\ D(t) &= K_p \cdot T_p \cdot \dot{G}(t) \end{split}$$
(b3)

The test reveals that impaired glucose tolerance humans have a defect in the second phase as earlier stated.

Other models are based on the measurement of c-peptide [Eaton 1980] that is secreted in equal amounts as insulin with advantages that c-peptide is hardly extracted by the liver and its kinetics are linear. A disadvantage, from a qualitative standpoint, of models based on c-peptide is that the measurement of c-peptide is worse than the measurement of plasma insulin concentrations. However, these models can be modified in such a way that it can be used with measurements of plasma insulin concentrations.

[Toffolo 1995] introduces a c-peptide minimal model for an IVGTT which was extended by Breda to an oral c-peptide minimal model [Breda 2001]. The model assumes that  $\beta$ -cell insulin secretion depends on glucose plasma concentrations (static component, delayed production of new insulin) and on the rate of change of glucose plasma (dynamic component, secretion of stored insulin). This model is comparable with Toffolo's model in 1980 that use the same static component, the same order (1<sup>sl</sup>) for clearance of insulin, but with no static delay and no dynamic component. The introduction of a dynamic component into the model of  $\beta$ -cell insulin secretion gives the model more physiological meaning. The model is shown in (b4, b5) where  $SR_s(t)$  [pM/min] is the static  $\beta$ -cell insulin response,  $SR_d(t)$  [pM] is the dynamic  $\beta$ -cell insulin response and  $1/\alpha$  [min] is the delay time in the static response.

$$SR(t) = SR_{s}(t) + SR_{d}(t)$$
  

$$\dot{SR}_{s}(t) = \alpha \{SR_{s}(t) - y \cdot [G(t) - h]\}$$

$$SR_{s}(0) = 0$$
(b5)
$$SR_{d}(t) = \begin{cases} \kappa_{D} \cdot \dot{G}(t) & \text{if } G(t) > 0\\ 0 & \text{if } G(t) \le 0 \end{cases}$$

[Breda 2002] test reveals that impaired glucose tolerance humans have a diminished  $S_I$ , but an unaltered static and dynamic component as Bergman [Bergman 1981] found in *obese* lower tolerance humans. This test also reveals a delay between  $\beta$ -cell insulin secretion and changes of glucose concentration in plasma. Several other papers show similar results, e.g. [Ehrmann 2002, Cretti 2001]. Other publications have been published having likewise models [Mari 2002].

Both Steil's model and Breda's model are compared in [Steil 2003] using Bergman's minimal model [Bergman 1979] during a CLAMP on healthy humans. Both models fit the  $\beta$ -cell insulin secretion profile with low residuals. Also, comparable parameters between both models are not statistical different. However, Steil's model has better performance than Breda's model, since Steil's model has parameters with lower variations, more stability in close loop systems and adjusts better insulin needs to varying  $S_1$  and endogenous glucose appearance. One drawback is that Steil's model is tested by a CLAMP and not during physiological circumstances like Breda's model that is tested during an OGTT. Since Steil's model represents very well the  $\beta$ -cell insulin secretion behaviour and has interesting closed loop properties for future application of the model into an automated portable continuous glucose monitoring system, this model is used.

## C.1 Insulin profiles

(min)	prof	ile 1'	Prof	ile 2 <sup>2</sup>
(11111)	I(t) [pM]	G(t) [mM]	I(t) [pM]	G(t) [mM]
-15	53.0	6.16	55.1	6.16
I	45.9	6.55	53.7	5.12
15	86.2	6.66	94.6	7.59
30	<sup>I</sup> 34	7.87	139	10.0
45	145	8.31	143	11.2
60	165	8.91	152	13.5
90	174	8.97	209	14.6
120	169	9.52	194	I4.4
180	196	8.86	133	9.52
240	162	8.91	100	4.18

Table c.1: Two insulin and glucose profiles [Wagenmakers 2006] measured during an OGTT (75 gram glucose)

## C.2 Samples Schedules

(min)	Sample schedules
I	0 1 2 3 4 5 6 7 8 9 10 240
2	0 10 20 25 30 35 40 45 50 55 60 65 70 80 90 120 180 240
3	0 15 30 45 60 90 120 180 240

Table c.2: Sample schedule 1, 2 and 3

### C.3 Datasets

Dataset I is generated using insulin profile I and sample schedule I. Dataset 2 is generated using insulin profile 2 and sample schedule 1. Dataset 3 is generated using insulin profile 1 and sample schedule 2. Dataset 4 is generated using insulin profile 2 and sample schedule 2. Dataset 5 is generated using insulin profile 1 and sample schedule 3. Dataset 6 is generated using insulin profile 2 and sample schedule 3.

Interstitial glucose  $\overline{G'(t)}$  is measured in all previous datasets by taking the mean over each 30 minutes by taking the first sample at the 30<sup>th</sup> minute and the last sample at 240<sup>th</sup> minute.

<sup>&</sup>lt;sup>1</sup> Subject no: 5 Visit 6 / Day 15 [Wagenmakers 2006] <sup>2</sup> Subject no: 8 Visit 6 / Day 15 [Wagenmakers 2006] <sup>3</sup> Conversion factor between [µU /l] to [pM]: 7.0625.

<sup>&</sup>lt;sup>4</sup> Conversion factor between [mg/dl] to [mM]: 1/180.16 1/mM

C.4 Simulated model

## Insulin profile 1



Figure c.1: Simulated glucose plasma, insulin plasma, interstitial glucose, interstitial insulin and the rate of appearance using profile 1, sample schedule 1 and true value shown in table 1

## Insulin profile 2



Figure c.2: Simulated glucose plasma, insulin plasma, interstitial glucose, interstitial insulin and the rate of appearance using profile 1, sample schedule 2 and true value shown in table 1

# Plots of model to model adjustment

Model 1



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Estimator: Gb

0.008













Figure D.1: Plots of model to model asjustment of model 1 using dataset 1 (appendix c) with no added white noise









Figure D.2: Plots of model to model asjustment of model 1 using dataset 2 (appendix c) with no added white noise





Figure D.3: Plots of model to model asjustment of model 2 using dataset 2 (appendix C) with no added white noise

# Model 3



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Figure D.4: Plots of model to model asjustment of model 3 using dataset 3 (appendix C) with no added white noise

## A priori parameter sensitivity for kabs

1/SNR	profile 1				profile 2			
(%)	Ib	Gb	beta	k	Ib	Gb	beta	k
0	40.0	72.5	99.4	103.7	31.4	59.1	145.0	143.3
5	40.0	40.0	131.7	130.7	17.2	58.9	184.0	166.9
ю	40.0	73.3	153.5	168.6	17.4	59.0	241.3	195.1

1/SNR		pro	file 1		profile 2			
(%)	p2	P5	S <sub>G</sub>	Si-liver	p2	P5	S <sub>G</sub>	Si-liver
0	12.7	0.829	0.720	1.10	2.39	0.818	1.40	1.18
5	5.36	0.657	0.381	3.03	9.44	3.04	0.0242	2.00
IO	3.86	3.33	2.42	4.46	9.93	1.60	1.38	91.6

## A priori parameter sensitivity for v

1/SNR		prof	ile 1		profile 2			
(%)	Ib	Gb	beta	k	Ib	Gb	beta	k
0	40.0	70.6	12.2	10.3	28.6	59.6	8.24	8.33
5	49.0	84.6	19.1	16.9	9.5	63.2	105.1	80.2
10	43.I	94.2	43.1	51.6	27.6	68.0	74.2	72.9

1/SNR		pro	file 1			рго	file 2	
(%)	p2	P5	S <sub>G</sub>	Si-liver	p2	P5	S <sub>G</sub>	Si-liver
0	6.12	I.00	0.903	0.328	1.81	1.36	2.19	1.16
5	20.3	3.21	3.29	4.42	9.47	20.8	2.11	10.4
IO	4.27	8.3	5.36	5.39	66.6	10.1	7.22	13.9

## A priori parameter sensitivity for p4

1/SNR		prof	file 1		profile 2			
(%)	Ib	Gb	Beta	k	Ib	Gb	beta	k
0	50.0	50.0	7.55	10.0	22.48	49.8	17.0	15.9
5	40.0	394.7	25.4	23.5	20.8	60.0	53.2	46.7
IO	40.0	63.0	39.0	41.9	56.7	307.6	52.63	157.2

i/SNR	profile 1				profile 2			
(%)	p2	P5	SG	Si-liver	p2	P5	S <sub>G</sub>	Si-liver
0	2.13	0.789	0.111	0.180	3.08	0.893	I.2I	1.19
5	6.89	3.32	2.53	3.84	6.96	2.58	I.49	10.86
10	4.35	8.78	3.10	8.77	108.4	8.63	5.84	60.86

A priori parameter sensitivity for Si-Glut4

1/SNR		pro	file 1		profile 2			
(%)	Ib	Gb	beta	k	Ib	Gb	beta	k
0	50.0	50.0	11.2	13.9	42.9	47.6	21.6	19.8
5	48.2	946.4	38.4	29.1	41.8	234.4	53.6	44.9
IO	48.7	770.4	52.4	48.1	56.7	307.6	52.63	157.2

1/SNR		pro	file 1			pro	file 2	
(%)	p2	P5	SG	Si-liver	p2	P5	SG	Si-liver
0	34.0	0.135	0.771	0.625	36.0	0.262	1.61	0.831
5	40.7	3.98	2.30	4.04	39.7	20.2	1.76	14.0
10	42.58	15.6	6.56	7.00	66.6	10.1	7.22	13.9

Table E.I: A priori parameter sensitivity for kabs, v, p4 and Si-glut4 using synthetic data and using two different insulin profiles (data sets 1, 2, respectively, **appendix c**) with added white noise having a 1/SNR of 0, 5 and 10%

# European Research activities On (Closed loop) artificial pancreas and automatic insulin delivery

## **Consulted sources:**

EMBase (ScienceDirect), MEDLINE (ScienceDirect), Elsevier BIOBASE (ScienceDirect), Pubmed, Pascal (Pica), Web of Science, Elsevier Scopus, Inspec (webspirs)

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## Matlab code

## **D.1 Basic files**

### D.1.1

### gluc\_param\_new.m

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end
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Gb = 0.055*(112+93)/2; [MM] balance are set as the the platma
GIb=3.38; %(mM) paseline aldered elecation in interstitium
end
if 0 - the post was Same
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rlasmi
Gb = 0.055*(165+169)/2; (MM) is the discontention is plasma Glb=3.38; [MM] beschive allowed on the straight num
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WARNAL LA CEMATIONAL CONTROLLARS THE LEFT FLORING FIRST OF MEXICITIE and Bieldy, alcol, Means, I. I. and ac-۰, Watalucci, S., Jort. Tr N. (1, 1, 1, 1, 1, 1, 1, 1, 1, Marzi, F. Morusan, sk. Bus office, s"Aksorption and Lodiscon on the solution of the boost of Lorando Lest," Sproceedings of the Sthemassical control togal contents of the «TEFF Ending of Medicine and Party , Sand, K(2005), S. 2005 k8 =3e-2; % [lxpM] the strength of the strength strength k -strength ~strength \ strength ~ strength ~ strength ~ strength ~  $p_2 = 0.113;$  Splemary rate of that  $l_{\rm ext} = l_{\rm ext}$  is represented in the object -Sachish DS = 75/10; (1), (1) measure (1) 1.20 production of all do by tactor <1710 € Kabs =2.89e-2; - provide the state of the st 307 First of the start family Similar emplying " lowest, a second to the sense court of the state betast Sthe opposite o cons.  $k = 0.014; \quad (m_1 + m_1 + m_2) = (m_1 + m_2) + (m_1 + m_$ V=2.497e-1; ) is made to the set of a set of volume per unit of buay sweight .(1.3860 - 11 a \* 18: 18 1 on) Stel post to be at  $p6 = 0.00763 \times 10;$  [1 to a late of a total structure of the plane of the decoder of the second structure of the second stru scaled %by factor 19 k7 = 0.063\*(1/0.21); if the second constraints are second constraints and the second constraints are second constraints ar Scompart ment SIcap = 2e-3\*0.0395\*(1/0.21)\*1e3; [Finite Lips] the set of even insular t o Stimulate insulation of sequences intervention by depillary are numeric scaled is taking let SILiver =  $1.11e^{-5*1e4*}(1/7.0625); + 1 min + LipM]$  the efficacy of Insulin to subject head of production and standard short we as taken by layer. concluding island les Siglut4 = 1.0e-2; [F or  $2 + 1.0e^{-1}$  or  $22 + 1.0e^{-1}$  consult to stimulate alucare with the theorem in the second secon interstitium Ascaled by factor 1 Coeff =3/75; [] wetto next between and of labeleed alucose divided stry the sam of labeled all is the con-

p\_true = [Ib, Gb, p2, DS, Kabs, Beta, k, V, p4, p5, p6, SIcap,... SIliver, SIglut4,Coeff]; if and is in the parameter space

```
SFixed initial conditions
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c_insulin1 = 7.0625*[(7.5+6.5)/2 12.2 19 20.5 23.3 24.6 23.9 27.7 23];
t_c_insulin1 = [t_insulin1' c_insulin1'];
(2) interstitual ableose data
t_insulin2 = [0 15 30 45 60 90 120 180 240];
c_insulin2 = 7.0625*[(7.5+6.5)/2 12.2 19 20.5 23.3 24.6 23.9 27.7 23];
t_c_insulin2 = [t_insulin2' c_insulin2'];
end
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c_insulin1 = 7.0625*[(6.4+4.5)/2 14.8 22.2 27 27.7 46.8 81 46 7.];
t_c_insulin1 = [t_insulin1' c_insulin1'];
32 interstitual glucose data.
t_insulin2 = [0 15 30 45 60 90 120 180 240];
c_insulin2 = 7.0625*[(6.4+4.5)/2 14.8 22.2 27 27.7 46.8 81 46 7.7];
t_c_insulin2 = [t_insulin2' c_insulin2'];
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t_insulin1 = [0 15 30 45 60 90 120 180 240];
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t_insulin2 = [0 15 30 45 60 90 120 180 240];
c_insulin2 = 7.0625*[(7.8 + 7.6)/2 13.4 19.7 20.2 21.5 29.6 27.5 18.8
14.2];
t_c_insulin2 = [t_insulin2' c_insulin2'];
end
```

5 squared 2 norm of the error time is residuals

```
a=1; %output wordhing factor of all manufactor concentration (1st output)
b=1; control wordhing factor factor factor fit all places concentration %(2nd | output)
c=1; control wordhing factor factor factor factor places concentration %(3nd | output)
sigma1=0.025; %variance of uncered relaxes all as concentration()s
output)
sigma2=0.05; %variance of interativel () all one (2nd output)
sigma3=1;%0.06; %variance of tabel relaxes of output () all one () all output)
```

&Plots

```
s make plots if plt :
if plt==1;
%Plots
figure; plot(t_c_insulin(:,1), t_c_insulin(:,2), ' '); hold
%Basefine level:
plot( [t_c_insulin(1,1) t_c_insulin(end,1)], [Ib Ib], '
r',' = sitth',1.5)
xlabel('fite tith',1.5)
xlabel('fite tith', '); ylabel('fite tith', '); ')
title('set tith', '); ')
end
```

### D.1.2

### gluc\_ode\_new.m

```
function dLout = gluc_ode_new(t,Lin,t_c_insulin,p)
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't_c_insulari
>p: [1b, 6b, p2, DS, Kabs, Bets, r, , r, j5, p6, Sleap,...
 Stilver, Stalut , Dettl; " If instances barameter space
soutputs function:
adiout: new state derivatives di(k) = [a'(k); dX(k); dY(k); dGqut(k)]
2
    column vector.
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ban.
idX = 3; "interstitut number of static [M], this column of vector
6.1.11
idY = 4; subtensional shows concentration[mM], fourth column of
vector Lan
idGgut = 5; the mass of flue set in the dust [1/kd], fifth column of
vector
 1.12
 fixer model parameters:
       p(1); %[pM] baseline inclusion inclusion in clasma
Ib =
       p(2); %[mM] baseline quark in relation in plasma
Gb =
° known/unkown model parameters.
```

Theose from blood plasma to provide the by capibles, connitment

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dLout = [dG; dGL; dX; dY; dGgut];

## D.1.3

### gluc\_sim\_new.m

```
function [c_glucose,c_glucose_labeled,c_interstitial_insulin,...
            c_interstitial_glucose,c_Ggut,c_Ggut_labeled]=gluc_sim_new (L0,...
            t_c_insulin,p,plt,TITLE)
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Sysode_options = odeset('kell s', b , 'scal i', t to);
    [t,L] = ode45(@gluc_ode_new,t_c_insulin(:,1),L0,ode_options,
    t_c_insulin,p);
```

```
xlabel('| | ')
Title('|
                             ');
tions+
subplot(322); plot(t_c_insulin(:,1), c_interstitial_insulin,
· · · )
Title(TITLE);
1∓qui ∈
Gb=p(2);
subplot(323); plot(t_c_insulin(:,1),c_glucose,' ', ' - - - - - - - + ',2);
hold ;
plot( [t_c_insulin(1,1) t_c_insulin(end,1)], [Gb Gb], '
',' : · · · · · · · · .',1.5)
-Baseline level:
· · · · )
Title(TITLE);
fiquio
subplot(324); plot(t_c_insulin(:,1),c_glucose_labeled,' ',
··',2);
xlabel('
                 ');
ylabel('
                                          •)
Title(TITLE);
time c
subplot(325); plot(t_c_insulin(:,1), c_interstitial_glucose,
xlabel(')
                '); ylabel...
(',
            i and
                                      ')
Title(TITLE);
Stiguie
subplot(326);
· )
Title(TITLE);
end
```

## D.1.4

### obj\_fn\_new.m

```
function error = obj_fn_new(p_var, p_fix, L0, data1, data2, data3,...
      t_c_insulin1,t_c_insulin2,a,b,c,sigma1,sigma2,sigma3,...
      t_exp,Nparameters,NoMean)
1.9%
Subjective function Maximum Lease Duras - Estimation of the Oral
Minimal
%Model of glurose kinetics.
5, 2, 3,
Withis function is used as opportive function for LSQNONLIN.
555
*Function:
We - op__in_Aw(p_var, p_dix, L , c__i),dat 2, data;t_c_insultal,...
St_c_Institud, a, b, , stamal, stamat, take, take, to we, NoAr andress)
tliputs function:
%p_var: parameter space which will read anated
*p_fix: parameter space which is 'ixes.
%L0: Fixed initial conditions, (1) (matthe durate concentration in
%plasma, LU(z) [pM] insults contend of an intervention,
$L0(3) - [mMs alocose concentration of anternant um, L0(4) - [1/kg]
solucose to, entrata a he an
Watal: data of unlabeled places (now ) from]
Sdata2: data of interstitual insuling [PM].
Wata3: data at laceled plasma plusses [mk].
St_C_insulin: [pM] [min] time and emplitude amples matrix of plasma
Sinsulin.1 - blood glacose franz interstitual flucose data.
Sa: output weighting thetor of las intent.
%b: output weighting factor ' - : cross
%c: output weighting factor ' - : cross.
%sigmal: variance of 1st output.
%sigma2: Variance of End output.
%sigma3: Valiance of Sno cutpet.
St_exp: sampleinterval [nit] of set.
SNparamters: number of dimensional parameter space p.
Sectimator: defines the parameter which is used to the model to model
Sugustment identifi alos.
Coutput function:
Serior: minimizate a variable () s. (ne of organ)
555
8
SHISTORY
Watch Up, Salvader Alvan, 1
                                      . 10 .
W29 Lan (14, Natil Set Rivel, 1
al' dun 03, Natal var Kiel, 1
<sup>6</sup> 3, €
senst inaction 14 difficultion and our and
                                     21 15
```

 $h_{\rm E}$  true = [1b, 4, pr. 25, Kato, Beta, S, V, PJ, Pb, pr. 5124p,...]

```
s SILiver, Siglut4, Coett];
```

### p\_estimating = [p\_fix(1),p\_fix(2), p\_fix(3), p\_fix(4), p\_var(1), ... p\_fix(5),p\_fix(6), p\_var(2),p\_var(3), p\_fix(7),... p\_fix(8), p\_fix(9), p\_fix(10), p\_var(4),p\_fix(11)];

%Simulating the Oral Minimal Model of duress kinetics using the estimating ill dimensional parameter space and heri terperates goodlin concentration . 'samples for datal.

plt=0; ``make No plots
[c\_glucose\_estimating,c\_glucose\_labeled\_estimating,...

```
c_interstitial_insulin_estimating,c_interstitial_glucose_estimating,...
c_Ggut_estimating)=gluc_sim_new(L0,t_c_insulin1,
p_estimating,plt,...
' ' ' ' ');
```

Simulating the Oral Minimal Model of glucose kinetics using the estimating of dimensional parameter graph and the interpolated insulin concentration complete to data?.

### if NoMean==0

```
> Interpolation data.
t_insulin2_interp1 = t_c_insulin2(1,1):1:t_c_insulin2(end,1); * [min]
c_insulin2_interp1=interp1(t_c_insulin2(:,1),t_c_insulin2(:,2),...
t_insulin2_interp1);
t_c_insulin2_interp1(:,1)=t_insulin2_interp1;
t_c_insulin2_interp1(:,2)=c_insulin2_interp1;
```

```
[c_glucose_estimating2_interp1,c_glucose_labeled_estimating2_interp1...
c_interstitial_insulin_estimating2_interp1,...
c_interstitial_glucose_estimating2_interp1,...
c_Ggut_estimating2_interp1]=gluc_sim_new(L0,t_c_insulin2_interp1,
...
```

```
p_estimating,plt, ' ');
```

```
* raivulating the mean of the interpolated data2 at the data2
samples
GIb=L0(4):
```

```
c_interstitial_glucose_estimating2_mean(1,1)=GIb;
i=2;
for r=t_c_insulin2(2:end,1)'
c_interstitial_glucose_estimating2_mean(i,1)=...
sum(c_interstitial_glucose_estimating2_interp1(r+1-t_exp:r))/t_exp;
i=i+1;
end
```
```
else
    [c_glucose_estimating2,c_glucose_labeled_estimating2,...
    c_interstitial_insulin_estimating2,...
    c_interstitial_glucose_estimating2,c_Ggut_estimating2]=...
    gluc_sim_new(L0,t_c_insulin2, p_estimating,plt,...
    '...+++');
end
```

```
end
```

```
slSQNOWLIN: objective function which instruct the memory encoded encoded
encoded encoded is a subset of the conservation of the encoded encoded
encoded encoded is a subset of the conservation of the encoded encoded
encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded encoded en
```

#### error3=c\_glucose\_labeled\_estimating-data3;

```
stotal model error
```

```
for the action of the arts and annous to tample, the time signs of
ATOTIO
Saud dat the set equal adding the store
since nativations
error1_new=error1;
error2_new=error2;
error3_new=error3;
data1_new=data1;
data2 new=data2;
data3_new=data3;
if length(error1)>length(error2)
   for g=1:(length(error1)-length(error2))
   error2_new(length(error2)+q)=0;
   data2_new(length(data2)+q)=1;
                                     of the creasing bryide by
Zert
  Sounds is line by: I show not have been a street since
win fioil www.
  Not errors, new is dready or
  end
   error1_new=error1;
  data1_new=data1;
  error3_new=error3;
  data3_new=data3;
else
tor q=1:(length(error2)-length(error1))
  error1_new(length(error1)+q)=0;
   data1_new(length(data1)+q)=1;
   error3_new(length(error3)+q)=0;
   data3_new(length(data3)+q)=1;
   end
```

```
error2_new=error2;
  data2_new=data2;
end
> is not the (a malitud best to be laceta. (i mazidataz_ae)).*.
.. enol_new; (...(namas.* acc_new)).*errors_new;
error=1e4*error1_new+1e4*error2_new+125*1e4*error3_new;
 alot muthe estimation process
1 f 0
figure(100)
subplot(411);
plot(t_c_insulin1(:,1),c_interstitial_insulin_estimating,'.');
xlabel(' = ' = ');
· ');
11 1 :
                                 drawnow;pause(.3);
3114036(100)
subplot(412);
plot(t_c_insulin1(:,1),c_glucose_estimating,' ',t_c_insulin1(:,1),...
   data1,' ');
xlabel('
ylabel('
               `);
                                   ');
                     ; ; . . .
Title({'!
  'E E E E E
                                  drawnow;pause(.3);
stiquie (106)
subplot(413);
plot(t_c_insulin1(:,1),c_glucose_labeled_estimating,' ',t_c_insulin1...
`);
                     ';...
Title({`
   '});
drawnow;pause(.3);
finer (100)
subplot(414);
if NoMean==0
plot(t_c_insulin2(:,1),c_interstitial_glucose_estimating2_mean, ' ',...
   t_c_insulin2(:,1),data2,' ');
else
plot(t_c_insulin2(:,1),c_interstitial_glucose_estimating2,':',...
```

t\_c\_insulin2(:,1),data2,' ');

end

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## D.2 model2model adjustment

### D.2.1

#### model2model\_bias\_new.m

```
0.00
Will Mibired M & L of Alwesse with a com-
SThis script identifies a posterior the Dral Minimal Model of glucose
kinetics using the model to model adjustment appreach for each a
A DR MALON
of the Letters and premove the
2.55
History
52 March 06, Salvador Almagro Frutos, TU/e
Similalizes all parameters of the PMM are all time samples of plasma
°insulin.
clear :;
NoMean=0;SCalculate the mean event of x conductor of interstitual allocs-
Scamples it set to C
gluc_param2_new;
Nparameters=length(p_true);
s adding boilse
data1_new=data1;
Adatal_new datal+0.025 * ranh(1) r b(us (1),1); $+6% RUIS
data2_new=data2;
data__aevenata242.05 * rank(lessb(a++),0;
data3_new=data3;
>data? prw-data(+1.)z! * can(b) rund() (..., b;
Synchor containing variation (\infty ) for the true 14
dimensional
sparameter space
rr=[-0.5 -0.45 -0.4 -0.35 -0.3 -0.25 -0.2 -0.15 -0.1 0 0.1 ....
   0.15 0.2 0.25 0.3 0.35 0.4 0.45 0.5];
Names of the ll simensional as unit a single
prosect[']b '; 'GE ', 'a '; 'D: ';...
'Kaba '; 'Esta '; 'k '; ' '; 'p: ';'p5
';....
    100
           '; 'Sloap ';'"Eliver';'Sliget';' weth '];
```

 $Model to model accustoment operated to each <math display="inline">\tau$  -imensions (P1,P2) of the

314-dimentional parameter solve(p).

%F1 p(1),F2 p(\*); 1: \*(1),F = (\*)... F1 p(1),F2 p(10) %P1 p(2),F2 p(\*); F1 b(-),F = (\*)... F1 p(2),F2 p(10) %P1 p(3),F2 p(\*); F1 p(3),F = (\*)... F1 p(\*),F2 p(10) % ... F1 p(\*),F2 p(\*); F1 p(3),F = (\*)... F1 p(\*),F2 p(10) % ... F1 p(\*),F2 p(\*);

```
estimator(1) =qq;
```

```
%linitialize the 200 estimation - calleter (Ed)
for yy=[5 8 9 13 14]%yy 4:(N) = orders (A)
bb=1;
    if yy>qq
```

estimator(2)=yy; contained real gety;

§ Saves all other parameters of the to domensional parameter set.

```
space.
```

```
p_fix=[];
for zz=1:Nparameters
    it zz~=estimator(1)
         if zz~=estimator(2)
         p_fix=[p_fix p_true(zz)];
         end
    end
end
```

for ii=1:length(rr)

```
for jj=1:length(rr)
```

```
%Fire incline of the . inmerional parameter
%space of the Star Minim | Model of aluresc kinetics
plt=0; %Make no plot.
iii=ii
jjj=jj
```

[resnorm]=gluc\_omm\_mls2\_new(data1, data2, data3, L0, t\_c\_insulin1, ...

```
t_c_insulin2,a,b,c,sigma1,sigma2,sigma3,p_true,p_fix,p_init,...
          t_exp,Nparameters,estimator,NoMean,plt);
          Ssaving the squared a nerror the error nurstree residual
          DataZ(ii,jj)=resnorm;
            Ssaving both true parameters of platting purpose (P1, P2)
            if rr(ii) == 0 && rr(jj) == 0
                resnormTrue=resnorm;
            end
             costileating assimily of the interest initial values.
            ıf ii==4 && jj==2 ||...
               ii==2 && jj==3 ||ii==4 && jj==8 || ...
               ii==15 && jj==7 || ii==6 && jj==11 ...
               || ii==8 && jj==3 || ii==12 && jj==5 || ii==3 && jj==14
|| ...
               ii==18 && jj==6 || ii==19 && jj==13 || ...
               ii==12 && jj==9 || ii==11 && jj==3
            clear :;
            [p_model,L0_model,c_glucose_model,c_glucose_labeled,...
c_interstitial_insulin_model,c_interstitial_glucose_model,...
c_Ggut_model,resnorm,residual,varp,stdp]=gluc_omm_mls_bias_new...
           (data1_new,data2_new,data3_new,L0,...
           t_c_insulin1,t_c_insulin2,a,b,c,sigma1,sigma2,sigma3,...
           p_fix,p_init,t_exp,Nparameters,estimator,NoMean,plt);
          ptrue1=p_true(qq) > n ally provided
ptrue2=p_true(yy) 
           estimation(bb,:)=[p_model(qq) p_model(yy)]
           bb=bb+1;
           end
          Reaving the indipation of the entry (1) the instance
          DataY(jj)=p_true(estimator(2))+rr(jj)*p_true(estimator(2));
          end
      Assuing the lud parameter (11) with rr(rr) +100% variation
     DataX(ii) = p_true (estimator(1)) + rr(ii) * p_true (estimator(1));
     end
     5 Makes 3d plots of the error tubetion (estimator), estimator?)
     "Contourlines 200;
    Minn=min(min(DataZ))
```

```
Maxx=max(max(DataZ))
    Contourlines=50
     ((min(min(DataZ))):0.01:(max(max(DataZ))))';
     if Contourlines>200
       Contourlines=200;
       display...
       ("Salka I aad
                                                     1
· · ·);
    end
    figure(333);
    contour(DataX,DataY,DataZ,Contourlines);
    hold ;
     a inservation for the couple of the free parameters as a reacted to
    plot(p_true(estimator(1)),p_true(estimator(2)),'s_');
    hold
    plot(estimation(:,1),estimation(:,2),' ');
    colormap per ( 15);
    colorbar;
    hold ;
    adaves plots of the error that is testimaterly estimators) in
119,
    Std.t etc.ep. tottot.
savename=[' ' pname(estimator(1),:) pname(estimator(2),:)
11.1 (a!]
    saveas(gcf, savename, '__') ___ trailingtingting', 'reuse',
'visible')
    savename=[' ' pname(estimator(1),:) pname(estimator(2),:)
'l.at.a']
    saveas(gcf, savename, ': ')
    savename=['' pname(estimator(1),:) pname(estimator(2),:)
11.04 ST
    saveas(gcf, savename, ' ')
    close(figure(100));
   end
   end
end
```

#### gluc\_param2\_new.m

1.2 Ronal Minimal Medel of glu. or. Kristin . į. sthis securit initializes all parameters of the Oral Minimal Model of Hucose skiperi s and the time sample of a least insuring 3History \$22 March 06, Salvador Almadro Frut 4, 19 4 329 Jan-04, Natal Van Rich, ED/c %17-Jup OB, Natal Yan Kird, THA 29% close i; - make plots if plt 1 plt=1; dixed model parameters: Walues from: 2002222 if 1 Subject po: 5 Visit of the is, is Ib = 7.0625\*(7.5+6.5)/2; [PM] because uselue conceptiation in plasma Gb = 0.055\*(112+119)/2; [mM] produce alweek conceptation in plasma GIb=3.38; [mM] caseline alweek transition in plasma ena if **0** % Subject ne: 15 - Visit ( ' bay 1'  $Ib = 7.0625 \star (6.4+4.5)/2; \times [pM]$  but lies insulin concentration in plasma Gb = 0.055\*(112+93)/2; [mM] baseline durate concentration in plasma GIb=3.38; s[mM] baseline qlucese succentration in interstitium enci if **0** % Subject no: 8 Visit 6 / 14V lt Ib = 7.0625\*(7.8 + 7.6)/2; N[pM] tastles involve consentration inpliama Gb = 0.055\*(165+169)/2; imm basiles decase concertration in plasma GIb=3.38; (mM) baseline ducos - sourcements in relition end \*Known unknown meter parameter (Mean)values obtained from nucles a correspondent) Annual International Conference : the IFLE Encomenting to Medicine and <sup>8</sup>Biology, Cancun, Mexice, 11-21 September 2013, ANatalucci, S., and F. Di Naido, P. S'afferani, C. De Marzi, P. Meresini, sk. Burattiri, "Absorption and Insulin Sensitive' than (in Aurory Televance Fest," Maccoordings of the 25th Annual Inc. Lational Conterpart t the

TEFE Et dage de les al les al les de v i trans a saution. DS = 75/10; f = 10; Lant Or \$1710 Kabs =2.89e-2; Hand proto an end that any approximation from the Jut and a block plasha Lit−t a s | sthe opposite de urs. k = 0.014; > [min] times distant lastric complying V=2.497e-1; , [loud] clasma above construction volume per date of body weight (1. Proc. 11 of 18 1. p5 =0.018;%[inadjusted cost of a large prenamed by lower into splood plasma. p6 = 0.00763\*10; (i modeles act of a state that the set to make scales by factor 10 k7 = 0.063 \* (1/0.21); [1 (1)] . pi smu alucose lite remote SCOMPATINESS SICap = 2e-3\*0.0395\*(1/0.21)\*1e3; 11 car l (2) h write etc. (1 usulta Lυ Solimplate insuling treasing on a second solution by empiricary Frequitment scaled by factor let SILiver = 1.11e-5\*1e4\*(1/7.0625); [loom \* LopM] the erroracy of insulin to Simulat negative production and structure of the optime by liver. Scaled by factor lea SIglut4 = 1.0e-2; & [1/min \* 1/pM] the effectively of rapidly to stimulate sqlucose clearance in muscle (vis (1-1) ) to the trans p4 =  $k7*k8; (11) \mbox{matrix} (14) \mbox{MJ}$  is the rate of plane shower inter-Univers struct a walled by fact to 11 Coeff =3/75; [] so the outer  $\omega$  is only for the loss above and the transformation of the loss of t p\_true = [Ib, Gb, p2, DS, Kabs, Beta, k, V, p4, p5, p6, SIcap,... SIliver, SIglut4, Coeff]; for the or of parameter space Placed initial onder his

$$\begin{split} & \text{LO}(1) = \text{GD}; \quad \text{f}[\text{mM}] \quad \text{diverses in a set of a s$$

\*Toput insuling conventiation is place (its); assumed to be known at - a th Simulation time sample from lieur et a letre of its measures. samples. "lime and amplitude samples of classes couldn: if 1 %Subject po: 5 Misit Marine 1 Sl > plood flucose data t\_insulin1 = [0 15 30 45 60 90 120 180 240]; c\_insulin1 = 7.0625\*[(7.5+6.5)/2 12.2 19 20.5 23.3 24.6 23.9 27.7 23]; t\_c\_insulin1 = [t\_insulin1' c\_insulin1']; 32 rinterstitial glucose data. t\_insulin2 = [0 15 30 45 60 90 120 180 240]; c\_insulin2 = 7.0625\*[(7.5+6.5)/2 12.2 19 20.5 23.3 24.6 23.9 27.7 23]; t\_c\_insulin2 = [t\_insulin2' c\_insulin2']; end if 0 % Subject no: 15 Pisit ( Fay 15 -1 - blood alucose data t\_insulin1 = [0 15 30 45 60 90 120 180 240]; c\_insulin1 = 7.0625\*[(6.4+4.5)/2 14.8 22.2 27 27.7 46.8 81 46 7.]; t\_c\_insulin1 = [t\_insulin1' c\_insulin1']; 32 Interstitial glucose data t\_insulin2 = [0 15 30 45 60 90 120 180 240]; c\_insulin2 = 7.0625\*[(6.4+4.5)/2 14.8 22.2 27 27.7 46.8 81 46 7.7]; t\_c\_insulin2 = [t\_insulin2' c\_insulin2']; end if 0 % Subject to : 6 Misit , ) <sup>r</sup> \*) blood alue se data t\_insulin1 = [0 15 30 45 60 90 120 180 240]; c\_insulin1 = 7.0625\*[(7.8 + 7.6)/2 13.4 19.7 20.2 21.5 29.6 27.5 18.8 14.2]; t\_c\_insulin1 = [t\_insulin1' c\_insulin1']; 2 leterstitul oheose at t\_insulin2 = [0 15 30 45 60 90 120 180 240]; c\_insulin2 = 7.0625\*[(7.8 + 7.6)/2 13.4 19.7 20.2 21.5 29.6 27.5 18.8 14.21; t\_c\_insulin2 = [t\_insulin2' c\_insulin2']; end 3 squared 2 norm of the error function residuals a=1; Soutput weighting factor of plasma glucose concentration (lst out put ) b=1; ) output weighting tick to the solutial physic concentration S(2nd output) c=1; 3 output weighting takes a time labeled during concentration \*(And output) sigmal=1; %variance of plasma (http://clum.ntration(ist output))

```
sigmal=1; & variance of integration of the due se (2nd output)
sigma3=1; & variance of plasma bare is the se (ind cutput)
```

```
%Plots
% make plots if plt 1
if plt==1;
figure; plot(t_c_insulin1(:,1), t_c_insulin1(:,2), ' '); hold
Sbaseline level:
plot( [t_c_insulin1(1,1) t_c_insulin1(end,1)], [Ib Ib], '
E',''L + WE = E',1.5)
xlabel('Lar (hora
                      1 I C
                                                                ')
                        · · · · )
title('A . . . .
end
Simulating the oral Miximal Skart is flores kinetics using the TKUE
%14-dimensional parameter space and the interpolated insulin
concentration
"samples for datal.
* sampleinterval (rus) of out (
t exp=15;
[c_glucose1, c_glucose_labeled1, c_interstitial_insulin1, ...
    c_interstitial_glucose1,c_Ggut1,c_Ggut_labeled1]=...
                                                           ');
    gluc_sim_new (L0,t_c_insulin1,p_true,plt,'
if NoMean==0
    % Interpolating difus
    t_insulin2_interp1 = t_c_insulin2(1,1):1:t_c_insulin2(end,1); t[nend]
    c_insulin2_interp1=interp1(t_c_insulin2(:,1),t_c_insulin2(:,2),...
    t_insulin2_interp1);
    t_c_insulin2_interp1(:,1)=t_insulin2_interp1;
    t_c_insulin2_interp1(:,2)=c_insulin2_interp1;
    S simulation bride islat
    [c_glucose_interpl,c_glucose_labeled_interp1,...
    c_interstitial_insulin_interpl,c_interstitial_glucose_interp1,...
    c_Ggut_interp1,c_Ggut_labeled_interp1]....
    =gluc_sim_new (L0,t_c_insulin2_interp1,p_true,plt,':. + + +');
    % calculating the mean of the interpolated data2 at the data2
samples
    GIb=L0(4);
    c_interstitial_glucose_mean(1,1)=GIb;
    i=2;
    for r=t_c_insulin2(2:end,1)'
    c_interstitial_glucose_mean(i,1) = ...
    sum(c_interstitial_glucose_interp1(r+1-t_exp:r))/t_exp;
    i=i+1;
    end
else
    ∛ SIMULATION USIL≮ dif i
    [c_glucose2, c_glucose_labeled2, c_interstitial_insulin2, ...
```

c\_interstitial\_glucose2,c\_Ggut2,c\_Ggut\_labeled2]....

```
102
```

```
=gluc_sim_new (L0,t_c_insulin2,p_true,plt,' ');
end
Sdata vector of the Oral Minimal Model of slucese kinetics using the
1 1 116-
If dimensional parameter space of the interpolated position
concentration.
camples.
Boatal: data of plasma algeose (eMu
data1=c_glucose1;
* sampleinterval [min] of data / cord conductor min_mis
t_exp=15;
Redata?: data of intenstitual diverse [nM]
1: NoMean==0
data2=c_interstitial_glucose_mean;
else
data2=c_interstitial_glucose2;
end
datas: data of plasma labeled place (PM)
data3=c_glucose_labeled1;
```

### gluc\_omm\_mls2\_new.m

```
tunction.
 [resnorm]=gluc_omm_mls2_new(data1, data2, data3, L0, t_c_insulin1, ...
                t_c_insulin2,a,b,c,sigma1,sigma2,sigma3,p,p_fix,p_init, t_exp, ...
                Nparameters, estimator, NoMean, plt)
 2 3 2
 SaloC_MM_PES2 Maximum salt groups from the transformed Minimal
Model
 %of glucose kinetics
 ų.
 3 6 3
 Efficient the estimates the property of the pr
 Statuetion of the Grad Mignal Model ' show a Ki other admitthe
 "Eparameters detried of all parameters
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 Star Charles
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 concentration,
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 Interstitial
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 $10(5) [likg] gladose concentration of the not.
 >t_s_insulum: [pM] [mun] fime and senditive complex matrix of plasma
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q \mapsto - \mathsf{whi}(h,\mathsf{will}) be estimated.
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&ROMean: It set to more, do not all late the mean over t<u>ext</u> minutes
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Sinterstitial alucese complex-
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%Outputs function:
Stesnorm: squared 2 norm of the cost of the cost.
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rew lo aimensional paramete, que
st: [1b, Sb, p2, DS, Kars, Bors, k, , bs, p*, is, Steip,...
Cliver, Slalust, settle
p_model = [];
offset=0;
for ww=1:Nparameters
    if ww==estimator(1) usean to the lot of analytic parameter and save it
        Sat the correct position . In the defined 14 dimensional
Larameter
        "space wetter (p.mcdel)
        offset=offset+1;
        p_model=[p_model p_init(1)];
        elseif ww==estimator(2) on the hermonic index parameter and 
Sparent of the end of a new of 14 dimensional
parameter
            separe concer (r_model).
        offset=offset+1;
        p_model=[p_model p_init(2)];
    else 3 save all the other present of at the correct position of
        314 dimensional parameter asa o (r_model)
        p_model=[p_model p_fix(ww-offset)];
    end
end
Sciential the Oral Maximal Mean of the science ( using the {\sf IRUF}
old dimensional parameter space and the interpolated inculing
concentiation.
'samples for datal.
[c_glucose1, c_glucose_labeled1, c_interstitia1_insulin1, ...
    c_interstitial_glucose1,c_Ggut1,c_Ggut_labeled1]....
    =gluc_sim_new(L0,t_c_insulin1,p_model,plt,'`'ill');
if NoMean==0
    Interpolating datas
    t_insulin2_interp1 = t_c_insulin2(1,1):1:t_c_insulin2(end,1);%[min]
    c_insulin2_interp1=interp1(t_c_insulin2(:,1),t_c_insulin2(:,2),...
    t_insulin2_interp1);
    t_c_insulin2_interp1(:,1)=t_insulin2_interp1;
    t_c_insulin2_interp1(:,2)=c_insulin2_interp1;
```

```
[c_glucose_interp1, c_glucose_labeled_interp1, ....
    c_interstitial_insulin_interp1, c_interstitial_glucose_interp1, ...
    c_Ggut_interp1,c_Ggut_labeled_interp1]....
    =gluc_sim_new (L0,t_c_insulin2_interp1,p_model,plt,'::::::'+');
    3 calculating the mean of the interpristed data2 at the data2
samples
    GIb=L0(4);
    c_interstitial_glucose_mean(1,1)=GIb;
    i=2;
    for r=t_c_insulin2(2:end,1)'
    c_interstitial_glucose_mean(i,1)=...
    sum(c_interstitial_glucose_interp1(r+1-t_exp:r))/t_exp;
    i = i + 1;
    end
else
    % simulation coincy datas
    [c_glucose2, c_glucose_labeled2, c_interstitial_insulin2, ...
    c_interstitial_glucose2,c_Ggut2,c_Ggut_labeled2]....
    =gluc_sim_new (L0,t_c_insulin2,p_model,plt,' ' ');
end
% error (function residuals)
error1=c_glucose1-data1;
⊥f NoMean==0
error2=c_interstitial_glucose_mean-data2;
else
error2=c_interstitial_glucose2-data2;
end
error3=c_glucose_labeled1-data3;
%total model error
For datasets have different concerpt simples, the unersions of
ACTORX.
hand datez its set equal many of
                                      t ce a
finitial values
error1_new=error1;
error2_new=error2;
error3_new=error3;
data1_new=data1;
data2 new=data2;
data3_new=data3;
it length(error1)>length(error2)
   for q=1:(length(error1)-length(error2))
   error2_new(length(error2)+q)=0;
  data2_new(length(data2)+q)=1; a not set to end of each vide by
1012
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knowns in the 121. If does not have menaline effect since
e.ietlinew
  set error2_new is already "er
   end
   error1_new=error1;
  data1_new=data1;
   error3_new=error3;
   data3_new=data3;
else
for q=1:(length(error2)-length(error1))
   error1_new(length(error1)+q)=0;
   data1_new(length(data1)+q)=1;
   error3_new(length(error3)+q)=0;
   data3_new(length(data3)+q)=1;
   end
   error2_new=error2;
  data2_new=data2;
end
% squared 2 norm of the error lunction residuals
resnorm1=1*(1/length(error1))*sum((error1_new).^2);
resnorm2=1*(1/length(error2))*sum((error2_new).^2);
resnorm3=125*1*(1/length(error3))*sum((error3_new).^2);
 stetal squared 2 norm of the error function residuals
```

```
resnorm=resnorm1+resnorm2+resnorm3;
```

## gluc\_omm\_mls\_bias\_new.m

function [p\_model,L0\_model,c\_glucose\_model,c\_glucose\_labeled\_model,... c\_interstitial\_insulin\_model,c\_interstitial\_glucose\_model,... c\_Ggut\_model,resnorm,residual,varp,stdp]=gluc\_omm\_mls\_bias\_new(data1,... data2,data3,L0,t\_c\_insulin1,t\_c\_insulin2,a,b,c,sigma1,sigma2,sigma3,... p\_fix,p\_init,t\_exp,Nparameters,estimator,NoMean,plt) 43. STOP\_MM\_MIS Maximum ( a top ) is constrained the oral Minimal Model 1 %glucose Kibeliež Yah #This turn test estimates the reader The Disk Prixible Service Schurters \*parameters action is also polyness. ÷. • . \*Function:  $\mathbb{E}[p_0] \mathbb{P}(de(1, \epsilon^{(1)}) \neq 1, 1, \epsilon^{(1)}) \mathbb{P}(e^{-\epsilon e_1 \pi - \theta e_1}, \epsilon^{(1)}) \mathbb{P}(e^{-\epsilon e_1}, \epsilon^{(1)}) \mathbb$ \_interverse distribute (meaning of the constant of the second data and .\_Gaut\_model,responded,earded,ear, the allocate and the second constructed, ... . data2,data3,CG,ty\_\_insulmi,ty\_\_inv= tr 2,a,; , , ional, (iona2, signa3,... p\_lix,p\_idil,t\_exp,Npataberet,websh,pil) aluputs tunction: Sdatal: data of intervitial model (pod %data2: data of remensional reactions Ydatas datas thatalas areas . . . . . No: Fixed for all in drack, seen administed a consistential in The is allowed through 1 \* \* - Catalog, (10.4) [10.4] "Slliver, Slifuda, Sedij; Podrbensi nalijarameter space a: caput weighting factor of the stput. <sup>3</sup> b: output weighting factor of 2nd output. sc: output weighting factor of Shit shipst. Ssigmal: Variance of 1st output. Ssignaz: Variable of Kid support. Sugmat: Variance of And output. "P\_11X: parameter space which in tikes. sp\_inu: initialism or stars as to and which will be estimated. Sampleinterval [Min] of anters. "Nparamers: ronder of annerol set counter space p. abstimator: defines the parameter which is used if the model to model

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          \left\{ \left\{ e_{i} \right\}_{i=1}^{\infty} \left\{ e_{i} \right\}_{i=1}
  1b = []; \quad (\uparrow ) (0 \le 0, 1 \ge 0, 1 \ge 1, 1); \quad (= 1)
 ub = []; here (x_1 + (x_1)); |];
          provides the the data acception shall be the option of the indi-
              lageratio.
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                        ',1e-20,... det mult: le l
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  [p_estimated, resnorm, residual, exitflag, output, lambda, jacobian]...
 =lsqnonlin(@obj_fn_bias_new,p_init,lb,ub,options,p_fix,L0,data1,data2,.
```

end

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YISUMODIAL Returns the months of a months
varp = resnorm*inv(jacobian'*jacobian)/length(t_c_insulin1(:,1));%
carlance
stdp = sqrt(diag(varp)); when states to vestige the equate post
Sof the Variance
% l' dimensional pseumetre spar-
p_model = [];
offset=0;
for ww=1:Nparameters
              if ww==estimator(1) from hour hour of material carameter and save it
                          %a the sources positive of the one detined 14 dimensional
parameter
                          %suade vector(p_model)
                          offset=offset+1;
                          p_model=[p_model p_estimated(1)];
                          elseif ww==estimator(2) in the high statute en and source that the intervention of the theory of the
Catal@eter
                                           55 - 1
                                                                                 (j + i 1)
                          offset=offset+1;
                          p_model=[p_model p_estimated(2)];
             else same di the Cherrison (e. 1) at the connect position of
%14 dimensional paramet ( para (c. 1))
                          p_model=[p_model p_fix(ww-offset)];
              end
end
L0_model = [p_fix(2), 0, 0, 0];
```

% make plots with new 14 dimensional parameter space if plt-1
[c\_glucose\_model,c\_glucose\_labeled\_model,c\_interstitial\_insulin\_model,.
..
c\_interstitial\_glucose\_model,c\_Ggut\_model]= gluc\_sim\_new (L0,...
t\_c\_insulin1,p\_model,plt,'\*
');

# obj\_fn\_bias\_new.m

```
tunction error = obj_fn_bias_new(p_var, p_fix, L0, data1, data2,
data3,...
                                   t_c_insulin1,t_c_insulin2,a,b,c,sigma1,sigma2,sigma3,...
                                    t_exp,Nparameters,estimator,NoMean)
pvar=p_var
   Observice tunorice Maximum react option. Extimation of the Dial
MI JMAI
    Morel + alberse kunstner.
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     Huberler:

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       t_______10301102, a, b, _______, sign (1, sign (2, s
      leputs function:
    4_200: parameter space which will be estimated
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    alueose esocontration i the e...
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    t_c_insulin: [pM] [and] Time and implicate samples matrix of plasma

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p_{1} (10, 11, 10, 12, 10, Ears, Ed., 5, 5, 61, p, 60, Sleup,...
Sluxer, Staluta;;
p estimating = [];
offset=0;
for ww=1:Nparameters
    parameter
       * Space Veet t(p Baar)
       offset=offset+1:
       p_estimating=[p_estimating p_var(1)];
       elsent ww==estimator(2) of the independent of parameter and
           Asave if at the structure of the dimensional
parameter
           address of the top Michelly
       offset=offset+1;
       p_estimating=[p_estimating p_var(2)];
    else %save all the other parameters as a the correct position of all dimensional automotors area (e.g. +1).
       p_estimating=[p_estimating p_fix(ww-offset)];
    end
end
Simulating the oral Minimal Model of and sole krastics to no the
estimatica.
the dimensional parameter grace red the arresponder is edim.
concentration.
%samples for datal.
plt=0; < make NO plots
[c_glucose_estimating,c_glucose_labeled_estimating,...
c_interstitial_insulin_estimating,c_interstitial_glucose_estimating,...
    c_Ggut_estimating]=gluc_sim_new(L0,t_c_insulin1,
p_estimating, plt, ...
    'h III.);
*Simulating the staf Maximal #soll of showing the
estimating
sl4 dimensional parameter space us the interpolated moduli
concentration
%samples for data2.
if NoMean==0
    S Hilmitel . Int . .
    t_insulin2_interp1 = t_c_insulin2(1,1):1:t_c_insulin2(end,1);*[num]
   c_insulin2_interp1=interp1(t_c_insulin2(:,1),t_c_insulin2(:,2),...
    t_insulin2_interp1);
```

```
t_c_insulin2_interp1(:,1)=t_insulin2_interp1;
t_c_insulin2_interp1(:,2)=c_insulin2_interp1;
```

```
[c_glucose_estimating2_interp1, c_glucose_labeled_estimating2_interp1...
    c_interstitial_insulin_estimating2_interp1,....
    c_interstitial_glucose_estimating2_interp1,...
    c_Gqut_estimating2_interp1]=gluc_sim_new(L0,t_c_insulin2_interp1,
. . .
    % calculating the mean of the interact conditions data? At the data?
samples.
   GIb=L0(4);
    c_interstitial_glucose_estimating2_mean(1,1)=GIb;
    i=2;
    for r=t c insulin2(2:end,1)'
    c_interstitial_glucose_estimating2_mean(i,1)=...
    sum(c_interstitial_glucose_estimating2_interp1(r+1-t_exp:r))/t_exp;
    i = i + 1:
   end
else
    [c_glucose_estimating2, c_glucose_labeled_estimating2, ...
    c_interstitial_insulin_estimating2,....
    c_interstitial_glucose_estimating2,c_Ggut_estimating2]=...
    gluc_sim_new(L0,t_c_insulin2, p_estimating,plt,...
      ······):
end
% SQNON'IN: objective function should return the model error:
"error: multimization variable of MLS (model error)
error1 = c_glucose_estimating-data1;
if NoMean==0
error2= c_interstitial_glucose_estimating2_mean-data2;
else SNo Mean
error2= c_interstitial_glucose_estimating2-data2;
```

```
error3=c_glucose_labeled_estimating-data3;
```

end

```
%total model error
> if datasets have different number of samules, the dimensions of
errorX
>add informate set equal address on
Statial alues
error1_new=error1;
error2_new=error2;
error3_new=error3;
data1_new=data1;
data2_new=data2;
data3_new=data3;
```

```
for q=1:(length(error1)-length(error2))
error2_new(length(error2)+q)=0;
```

```
·ero
             social in the 127. It for this sector effort since
 errorl_new
           action 2_DOW & Alterity of
            erd
            error1_new=error1;
            data1_new=data1;
             error3_new=error3;
            data3_new=data3;
 else
  for q=1:(length(error2)-length(error1))
            error1_new(length(error1)+q)=0;
             data1_new(length(data1)+q)=1;
             error3_new(length(error3)+q)=0;
            data3_new(length(data3)+q)=1;
            end
             error2_new=error2;
            data2_new=data2;
 end
 setror:(a./(sigmal.*datil_new)).******i_new(B. (sigmax.*data2_new)).*.
  . .
                  error2_new+(c./(stamas.*datas_new)).*error3_new;
    30
 error=le4*error1+1*le4*error2+l25*le4*error3;
 Serror-(a./(Signal *total pow)). *** at new;
 Service (b./(Signal Addition ow)).
 Series (C./ Grander Martin (Sev)). Series as we see the second se
 * plots our namest matrix resource
 if 1
 figure(100)
 subplot(411);
plot(t_c_insulin1(:,1),c_interstitial_insulin_estimating,':');
 xlabel(''__Be___be_et');
');
Title({'h (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (1997) (19977) (19977) (19977) (19977) (19977) (19977) (19977) (19977
                                                                                                                                                 drawnow;pause(.3);
 stigure(190)
 subplot(412);
 plot(t_c_insulin1(:,1),c_glucose_estimating,' ',t_c_insulin1(:,1),...
                 data1,' ');
 xlabel(''lime = bill(');
');
                                                                                                                                                                                                                                  '});
drawnow;pause(.3);
^{\circ} ( Lyper-(106)
 subplot(413);
plot(t_c_insulin1(:,1),c_glucose_labeled_estimating,':',...
```

```
t_c_insulinl(:,l),data3, ' ');
```

```
·);
                                    drawnow;pause(.3);
Ptionre(100)
subplot(414);
if NoMean==0
plot(t_c_insulin2(:,1),c_interstitial_glucose_estimating2_mean,'.',...
 t_c_insulin2(:,1),data2,' ');
else
plot(t_c_insulin2(:,1),c_interstitial_glucose_estimating2,',',...
t_c_insulin2(:,1),data2,' ');
end
xlabel(' = [ • ');
ylabel(' = ' • ';..
Title({' ;...
                                             :·· '});
drawnow;pause(.3);
e nd
```

# D.3 worstcase bias

### worstcase\_bias\_ss.m

```
tunction [WorstCase_Bias, estimation] = worstcase_bias_SS()
```

```
5 5 6
Woral Minimal Model of glucose Kinethes.
ò
aThis script calculates the wordthank of the stal Manual Model of
alucos
 5
$ 6 %
&distory
$25-08-06, Salvador Almagro Frutos, 1976
もこも
ちもな
%Function:
&[WorstLase_Bras] weistcase_bla ()
6
&Input: functi 4:
5 NONE
Soutputs function:
Worstlasejs in which is a reaction of the Restlicable watalance (is a state of the Restlicable state) and the Restlicable state of the Restlicable
sparameters of a state, , ,
7 InitiAlisati e
clear ii
gluc_param_new;
Nparameters=length(p_true);
NoMean=0;%Calculates the mean over the monster of interstitual
glucose
%samples if NoMean 0
diary [1] \approx \alpha , sets the names of the nativaled parameters:
filename='
                                                                                                                                              ÷;
Sp: (16, 36, pz, 3, kubb, Beth, s_{i} , ps, p_{i}, k_{i} , other, ...
"Stliver, Staluti, "ortil: 15 gmen achal presmoter state
ъв кноми разалестия:
p_fix = [Ib Gb p2 DS Beta k p5 p6 SIcap SIliver Coeff];
% using optimal incom
DS=75/10; [bd_l_optlopt_optlop
36 unknown model parageter:
p_init = [Kabs V p4 SIglut4];
```

1\_\_\_\_\_

a cample scheenles with a frequency of a musule for data to

```
t_insulin1_interp1 = t_c_insulin1(1,1):1:t_c_insulin1(end,1); <[min]
c_insulin1_interp1=interp1(t_c_insulin1(:,1),t_c_insulin1(:,2),...
    t_insulin1_interp1);
t_c_insulin1_interp1(:,1)=t_insulin1_interp1;
t_c_insulin1_interp1(:,2)=c_insulin1_interp1;
t_insulin1 =1:1:t_c_insulin1(end,1)
St_insulial -[0.10.20.20.30 20 if if is 50.55 60.65 20.80.90.120.180.240];
**_insulin1 [0.15/30/45/60/30] (0.180/2001;
tor uuu=1:length(t_insulin1)
c_insulin1(uuu) = c_insulin1_interp1(t_insulin1(uuu)+1);
end
clear
t_c_insulin1(:,1)=t_insulin1;
t_c_insulin1(:,2)=c_insulin1;
% simulation using data1
[c_glucose_model,c_glucose_labeled_model,...
    c_interstitial_insulin_model,....
    c_interstitial_glucose_model,c_Ggut_model]=...
    gluc_sim_new(L0,t_c_insulin1, p_true,plt,...
         +1 + + + + ');
if NoMean==0
    S Diterpolating datas
    t_insulin2_interp1 = t_c_insulin2(1,1):1:t_c_insulin2(end,1); [mi]
    c_insulin2_interp1=interp1(t_c_insulin2(:,1),t_c_insulin2(:,2),...
    t_insulin2_interp1);
    t_c_insulin2_interp1(:,1)=t_insulin2_interp1;
    t_c_insulin2_interp1(:,2)=c_insulin2_interp1;
    * special processing and as
    [c_glucose_model2_interp1,c_glucose_labeled_model2_interp1,...
    c_interstitial_insulin_model2_interp1,...
    c_interstitial_glucose_model2_interp1, ...
    c_Ggut_model2_interp1,c_Ggut_model2_labeled_interp1]....
    \pm calculating the mean of the interpedated datas of the datas.
samples
    GIb=LO(4):
    c_interstitial_glucose_mean(1,1)=GIb;
    i=2;
    for r=t_c_insulin2(2:end,1)'
    c_interstitial_glucose_mean(i,1) = ...
    sum(c_interstitial_glucose_model2_interp1(r+1-t_exp:r))/t_exp;
    i=i+1;
    end
else
    % simulation using data`
    [c_glucose_model2,c_glucose_labeled_model2,...
       c_interstitial_insulin_model2, ...
```

```
c_interstitial_glucose_model2, c_Ggut_model2, c_Ggut_labeled_model2]...
    end
Sladding in ite
datat_new=t_insulin1;
data1_new=c_glucose_model;
stated new details according to the contract of the call sewight; due RULD
if NoMean==0
data2_new=c_interstitial_glucose_mean;
.deta2_new data2_new+0_10 * code(1 = chad t agreew), 1); .+8 _ KUIS
else
data2_new=c_interstitial_glucose_model2;
Suata2_new acta2_new(0.10 * readn(looth(ista2_new),1); 508% RUIS%
end
sdata3_new_awgn(c_gturess_lubelessim_scl interpl,(1/0.)6),'measured',...
 % 'linear'); erec RELS: inflore direc
data3_new=c_glucose_labeled_model;
s end initiali ati m
&Capilating the determinant of the true contractor variations of the
Aidentifiathe scorebor by the light of the end off
p_init_copy=p_init;
b=1;
estimation(1, :) = [-0.5 - 0.4 - 0.3 - 0.2 - 0.1 0.1 0.2 0.3 0.4 0.5];
estimation1(1,:)=estimation(1,:);
estimation2(1,:)=estimation(1,:);
estimation3(1,:)=estimation(1,:);
estimation4(1,:)=estimation(1,:);
for kk=1:length(p init)
    11=0;
    for zz=estimation(1,:);
       p_init=p_init_copy;
       11=11+1;
        estimator=find(p_init(kk)==p_true);
       estimator1=find(p_init(1)==p_true);
       estimator2=find(p_init(2)==p_true);
       estimator3=find(p_init(3)==p_true);
       estimator4=find(p_init(4)==p_true);
        Svarying the identiable parameters by 22
       p_True=p_true(estimator) and a Lay in monitor
       p_init(kk)=p_init(kk)+ zz*p_init(kk)
        [p_model,L0_model,c_glucose_model,c_glucose_labeled_model,....
       c_interstitial_insulin_model,...
       c_interstitial_glucose_model,
c_Ggut_model, resnorm, residual, varp...
        ,stdp]=gluc_omm_mls_new(data1_new,data2_new,data3_new,L0,...
        t_c_insulin1,t_c_insulin2,a,...
b,c,sigma1,sigma2,sigma3,p_fix,p_init,t_exp,Nparameters,NoMean,plt);
       estimator
```

```
estimation(kk+1,11)=p_model(estimator);
                         estimationl(kk+1,11)=p_model(estimatorl);
                         estimation2(kk+1,11)=p_model(estimator2);
                         estimation3(kk+1,ll)=p_model(estimator3);
                         estimation4(kk+1,ll)=p_model(estimator4);
                        end
                         Calculating worr filse
                         estimation
                         WorstCase_Bias(b,:) = [p_true(estimator) ...
                         max(abs(p_true(estimator)-estimation(kk+1,:)))];
                         b=b+1;
                         end
                         TARGET=fopen(filename, ' ');
                          fprintf(TARGET,...
                                                                                                                                                                           · · · · ·
1 . . .
                                                                                                          ',estimation1');
                        fprintf(TARGET,...

    1.1. And the state of the second secon
                                                                                                           ',estimation2');
                         fprintf(TARGET,...
                          10.101 10.101
                                                                                                           ',estimation3');
                         fprintf(TARGET,...
                          1.1. 16.101
                                                                                                                 .1 1.1 +.1
                                                                                                          ',estimation4');
```

```
fclose(TARGET);
```

# D.4 A priori parameter sensitivity

#### param\_sens\_new.m

tunction [varp\_matrix]=param\_sens\_new() ちもも \*Oral Minimal Model of glucose Eric in 1001 STALE FINIAL FRANCE A process of the second state of the second st 法贸易 ž ※History \$07-july 06, Calvidor Almadro Frutis, 1070 \* % % %Function: %[Varp\_matr X]=param\_seus\_new() 5 \*Inputs function: . NONE ŝ, "Outputs function: Svarp matrix: beterminant of the orbit of the orbits for variations of the a Spriori parameters by S., S., L., J., C. on Sta

> initialization
clear if
gluc\_param\_new;
Nparameters=length(p\_true);
NoMean=0;%Calculates the mean over f\_exp minutes of interstitial
glucose
%samples if NoMean 0

```
%p. [1b, Gb, b2, b5, Kabo, Peta, K, , pa, p<sup>4</sup>, pb, Sloap,...
%pliver, Slalut4, 'oett];% 1' dimensioned parameter space
%% known parameters:
p_fix = [Ib Gb p2 DS Beta k p5 p6 SIcap SIliver Coeff];
% using optimal input
DS=75/10;%[DS_T_opt+ optimal(t))
%% utknown model parameter :
p_init = [1.4*Kabs 1.4*V 1.5*p4 1.5*SIglut4];
```

s sample schedules with a trapped start number for data 1
t\_insulin1\_interp1 = t\_c\_insulin1(1,1):1:t\_c\_insulin1(end,1);%[min]

```
c_insulin1_interp1=interp1(t_c_insulin1(:,1),t_c_insulin1(:,2),...
    t_insulin1_interp1);
t_c_insulin1_interpl(:,1)=t_insulin1_interpl;
t_c_insulin1_interpl(:,2)=c_insulin1_interpl;
 simulation using data1
[c_glucose_model_interp1,c_glucose_labeled_model_interp1,...
    c_interstitial_insulin_model_interp1,....
    c_interstitial_glucose_model_interp1,c_Ggut_model_interp1]=...
    gluc_sim_new(L0,t_c_insulin1_interp1, p_true,plt,...
                 ');
if NoMean==0
    5 Interpolating data2
    t_insulin2_interp1 = t_c_insulin2(1,1):1:t_c_insulin2(end,1);%[min]
   c_insulin2_interp1=interp1(t_c_insulin2(:,1),t_c_insulin2(:,2),...
   t_insulin2_interp1);
    t_c_insulin2_interp1(:,1)=t_insulin2_interp1;
   t_c_insulin2_interp1(:,2)=c_insulin2_interp1;
    3 simulation using data2
    [c_glucose_model2_interp1, c_glucose_labeled_model2_interp1, ...
    c_interstitial_insulin_model2_interp1,...
   c_interstitial_glucose_model2_interp1,...
   c_Ggut_model2_interp1,c_Ggut_model2_labeled_interp1]....
    3 calculation the mean of the integral lated datas of the data?
samples
   GIb=L0(4);
   c_interstitial_glucose_mean(1,1)=GIb;
   i=2;
   for r=t_c_insulin2(2:end,1)'
   c_interstitial_glucose_mean(i,1) = ...
    sum(c_interstitial_glucose_model2_interp1(r+l-t_exp:r))/t_exp;
   i=i+1;
   end
else
    ' simulation using satir
    [c_glucose_model2, c_glucose_labeled_model2, ...
       c_interstitial_insulin_model2,...
c_interstitial_glucose_model2,c_Ggut_model2,c_Ggut_labeled_model2]...
   =gluc_sim_new (L0,t_c_insulin2,p_true,plt,'' ');
end
 adding monae
datat_new=t_insulin1_interp1;
data1_new=c_glucose_model_interp1;
data1_new=data1_new+0.05 * randn(length(data1_new),1); 3+6% RUIS
if NoMean==0
data2_new=c_interstitial_glucose_mean;
data2_new=data2_new+0.10 * randn(length(data2_new),1); + 9 + P**IS
```

```
else
data2_new=c_interstitial_glucose_model2;
data2_new=data2_new+0.10 * randn(length(data2_new),1); >+3% KULS%
end
data3 new=c glucose labeled model interpl;
8 end int talization
staculating the determinant of the traded matrix for variations of the
%a priori parameters by 30, 20, 10, (, 1, 20 and 30%.
p_fix_copy=p_fix;
estimation(1,:)=[-0.3 -0.20 -0.1 -0.05 0.05 0.1 0.20 0.30];
estimation1(1,:)=estimation(1,:);
estimation2(1,:) = estimation(1,:);
estimation3(1,:)=estimation(1,:);
estimation4(1,:)=estimation(1,:);
for kk=1:length(p_fix)
   11=0;
    p_fix=p_fix_copy;
    for zz=estimation(1,:)
        p_init_saved = [Kabs V p4 SIglut4];
        estimator1=find(p_init_saved(1)==p_true)
        estimator2=find(p_init_saved(2)==p_true)
        estimator3=find(p_init_saved(3)==p_true)
        estimator4=find(p_init_saved(4)==p_true)
        11=11+1;
        &valying the a criori parameters by sc*1000
        p_fix(kk) = p_fix(kk) + zz*p_fix(kk)
        [p_model,L0_model,c_glucose_model,c_glucose_labeled_model,....
            c_interstitial_insulin_model,...
        c_interstitial_glucose_model,
c_Gqut_model, resnorm, residual, varp...
        ,stdp]=gluc_omm_mls_new(data1_new,data2_new,data3_new,L0,...
t_c_insulin1_interp1,t_c_insulin2,a,b,c,sigma1,sigma2,sigma3,...
        p_fix,p_init,t_exp,Nparameters,NoMean,plt);
        estimation1(kk+1,ll)=p_model(estimator1);
        estimation2(kk+1,ll)=p_model(estimator2);
        estimation3(kk+1,ll) =p_model(estimator3);
        estimation4(kk+1,ll)=p_model(estimator4);
        varp_matrix(kk,ll)=det(varp) = atstate, in monition
    end
end
        TARGET=fopen(filename,
                                 );
        fprintf(TARGET,...
              ',estimation1');
 1 \rightarrow 1 = 1
        fprintf(TARGET,...
```

'! ... '...' fprintf(TARGET,... '...' fprintf(TARGET,...' '...' fprintf(TARGET,... '...' fclose(TARGET); varp\_matrix=[p\_fix' varp\_matrix]

# **D.5 Optimal Input**

### opt\_input\_new.m

```
function [DS_T_opt,DS_matrix]=opt_input_new()
788
SOral Minimal Model of Hucebe Erge rule t
Alfris functions of balance the screense open on the scale Menned Model
Set ghave during the accordance to the action of the
х н <sub>10</sub>
Istory
107 july 06, Salvador Almadro Frat..., 10 e
olo
Olo
Shunction:
%[DS_T_opt,DS_matrix] opt_input new 0
 Inputs Conctron:
    NONE
Poutputs tunetion.
SDS\_T\_+p optimal input.
SDS_BRATEIX: BALEIX - TORRED - LET - XF OR SET CHARGENE OF FISHER BALLEX
S(and tow)
& indiates to at each
gluc_param_new;
Nparameters=length(p_true);
NoMean=0; staleulate the mean over the spectrum deposit interprintial glucose
∜samples
plt=0; anake o geter
% Sets the name. of the name to a construct the estimated parameters:
filename='() = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0) = (0,0
sp: - [10, 6b, p2, 58, Kabs, b-ta, 2, p4, p5, p6, Sleap,...
«Siliver, Sights, Court]: The norse of parameter space
sknown parameters:
p_fix = [Ib Gb p2 DS Beta k p5 p6 SIcap SIliver Coeff];
 SunKnows, model parameters.
p_init = [1.4*Kabs 1.4*V l.5*p4 1.5*SIglut4];
3 sample schedules with a frequency of 1 monute for data 1
t_insulin1_interp1 = t_c_insulin1(1,1):1:t_c_insulin1(end,1); ([min]
c_insulin1_interp1=interp1(t_c_insulin1(:,1),t_c_insulin1(:,2),...
        t_insulin1_interp1);
t_c_insulin1_interp1(:,1)=t_insulin1_interp1;
t_c_insulin1_interp1(:,2)=c_insulin1_interp1;
      simple whereast with the average
                                                                                             a construction data as
  t_insulin2_interp1 = t_c_insulin2(1,1):1:t_c_insulin2(end,1); ([min])
  c_insulin2_interp1=interp1(t_c_insulin2(:,1),t_c_insulin2(:,2),...
```

```
t_insulin2_interp1);
 t_c_insulin2_interp1(:,1)=t_insulin2_interp1;
 t_c_insulin2_interp1(:,2)=c_insulin2_interp1;
3 end initialization
9 different values for 0. (Amation 400 + and DG)
DS_vector=[1:100]./10; male by fermile (see differential
equations)
kkz=0;
for z=1:length(DS_vector)
    kkz=kkz+1
    p_init_saved = [Kabs V p4 SIglut4];
    estimator1=find(p_init_saved(1)==p_true)
    estimator2=find(p_init_saved(2)==p_true)
    estimator3=find(p_init_saved(3)==p_true)
    estimator4=find(p_init_saved(4)==p_true)
  charge PS
DS=DS_vector(z)
p_true(4)=DS;
p_fix = [Ib Gb p2 DS Beta k p5 p6 SIcap SIliver Coeff];
 imulating models, india sample schemule with a trequency of 1 minute.
  nor plana (t. p. ).
[c_glucose_model_interp1, c_glucose_labeled_model_interp1, ...
    c_interstitial_insulin_model_interp1, ....
    c_interstitial_glucose_model_interp1,c_Ggut_model_interp1]=...
    gluc_sim_new(L0,t_c_insulin1_interp1, p_true,plt,...
     11 - 11 - L + ');
ií NoMean==0
    ? Interpolation data?
    % simulation using data2
    [c_glucose_model2_interp1,c_glucose_labeled_model2_interp1,...
    c_interstitial_insulin_model2_interp1, ...
    c_interstitial_glucose_model2_interp1,...
    c_Ggut_model2_interp1,c_Ggut_model2_labeled_interp1]....
    =gluc_sim_new (L0,t_c_insulin2_interp1,p_true,plt,' == (L0,t_1,');
    A calculating the mean of the processing datas at the datas
samples
    GIb=L0(4);
    c_interstitial_glucose_mean(1,1)=GIb;
    i=2;
    for r=t_c_insulin2(2:end,1)'
    c_interstitial_glucose_mean(i,1) = ...
    sum(c_interstitial_glucose_model2_interp1(r+1-t_exp:r))/t_exp;
```

```
i=i+1;
    end
else
    a simulation using lature
    [c_glucose_model2, c_glucose_labeled_model2, ...
        c_interstitial_insulin_model2,...
c_interstitial_glucose_model2,c_Ggut_model2,c_Ggut_labeled_model2]...
    =gluc_sim_new (L0,t_c_insulin2,p_true,plt,'
                                                          ');
end
6 adding or re-
datat_new=t_insulinl_interp1;
data1_new=c_glucose_model_interp1;
data1_new=data1_new+0.025 * randn(length(data1_new),1);
if NoMean==0
data2_new=c_interstitial_glucose_mean;
data2_new=data2_new+0.05 * randn(length(data2_new),1);
else
data2_new=c_interstitial_glucose_model2;
data2_new=data2_new+0.05 * randn(length(data2_new),1);
end
data3_new=c_glucose_labeled_model_interp1;
"splots of data with a real to a c
* make plata it prt 1
if plt==1
%Plots
figure;
Gb=p_true(2);
hold on;
plot(t_c_insulin1_interp1(:,1),c_qlucose_model_interp1,' '',
'Linewicth',2);
hold on;
plot( [datat_new(1) datat_new(end)], [Gb Gb], ' - ',' - ..., ', 1.5)
%Baseline level:
xlabel('ii | . ]
                    ');
                                                    • )
ylabel('
Title(...
11 . 1
         the action of the
                                              4 (1) (1) (1) (1) (1)
(ashuan');
stigur⊬
if NoMean==0
subplot(312); plot(t_c_insulin2(:,1), data2_new,' ', ' + .... + ...',2)
hold .
plot(t_c_insulin2(:,1),c_interstitial_glucose_mean,' ::',...
    ''.itewisth',2);
else
subplot(322); plot(t_c_insulin2(:,1), data2_new,' ', ' + ..., +',2)
hold
plot(t_c_insulin2(:,1),c_interstitial_glucose_model2,' · ',...
     стери (11',2);
end
```

```
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```

```
')
 ('T i i
Title(...
              1, 1 | 1
                                       b i i
· · · · · ·
( ( ) ( ) ( ) ;
hold ;
plot(t_c_insulin1_interp1(:,1),c_glucose_labeled_model_interp1,' ',
····
/ 1 1 1 1 1 1 / ( , 2) ;
    ;
hold
Baseline lovel:
xlabel(
                  ');
                                 · · · · · )
ylabel(' I
Title(...
, the set of the
 · · ');
end
"estimating the model parameters.
[p_model,L0_model,c_glucose_model,c_glucose_labeled_model,...
c_interstitial_insulin_model, ...
c_interstitial_glucose_model,c_Ggut_model,resnorm,residual,varp,...
stdp]=gluc_omm_mls_new(data1_new,data2_new,data3_new,L0,...
t_c_insulin1_interp1,t_c_insulin2,a,b,c,sigma1,sigma2,sigma3,p_fix,...
p_init,t_exp,Nparameters,NoMean,plt);
  )f z==1
    varp_old=varp 1 cannot be a first her matrix
    DS_T_opt=DS_vector(z)
  elseit det (varp) < det (varp_old) ' ' i i the tide metrix
   varp_old=varp; avoid the time of action name of the timber matrix
   det (varp) la sela sessent a
   DS_T_opt=DS_vector(z)
  end
  det (varp)
  DS_matrix(z,:)=[DS,det(varp)];
      estimationl(kkz)=p_model(estimator1);
      estimation2(kkz)=p_model(estimator2);
      estimation3(kkz)=p_model(estimator3);
      estimation4(kkz)=p_model(estimator4);
end
      T.
```

TARGET=fopen(filename, '	');
fprintf(TARGET, '	<pre>,estimation1');</pre>
fprintf(TARGET,'	',estimation2');
fprintf(TARGET, ' 🕛 .	<pre>,estimation3');</pre>
fprintf(TARGET, '	',estimation4');
fclose(TARGET);	

## D.6 Optimal Sample Schedule

## opt\_sample\_schedule\_new.m

```
function [opt_SS_data1, opt_SS_data2, opt_data1, opt_data2...
     , opt_data3]=opt_sample_schedule_new()
Strat Marthae Marte
                          - 1142 - Arrison - 1980
This for tarm tells f , the process in president (88) for the
and Minimal Model at alumnation and the Raterin act of the fisher
matrix.
sAlgorithms cased on s aring with conventional sample schedule where
weach individual sample is optimal try tomparine it with its
neighbouring
stamples in the conventional samples chemic.
o.
. 100
U
SHISTORY
507 Buly D6, Silvidor Almonis Fills, 1996
, N. 8
stunction:
lopt_SS_data1, opt_ui_outur, j = u = d = , off_data1, opt_fite...
     ,opt_data() opt_sample_schedule_cw()
э
S.
slnputs function:
5 NONE
4
Soutputs function:
opt_SS_datal: octime sample cherale for an i.
sopt_ss_datar: oction sample inclusion of inclusion
sopt_ss_datar: of inclusion in the oction of the
sopt_datal: of inclusion to the oction of the oction
sopt_datar: of inclusion to strength of the oction of the oction
sopt_datar: of inclusion to strength of the oction.
% initialisation
clear +1
NoMean=0;%Calculates the mean over to spontates of interstitual
alucose
Max_Exp_Dur=240; Its maximal abec for the experiment cun take
Ssamples if NoMean-0
gluc_param2_new;
Nparameters=length(p_true);
                                                    il L. Z.
diary a course of the
```

Spi = [10, 60, p., 15, Kabs, Bets, c. 1, pe, ph, pe, Slippin, Siliver, Slipte, certip, is done a ball-barateter space

3 using optimal input
```
DS=75/10; * ["S_[_opt] root_input()
```

```
CREARING PERMINATOR ST
A TIX HE BOARD, NAME AND A DE
p_fix = [Ib Gb p2 DS Beta k p5 p6 SIcap SIliver Coeff];
initial value unenews model parameters:
r_mut [p4, p5, p6, along, firs , [matt];
p_init = [1.4*Kabs 1.4*V 1.4*p4 1.4*SIglut4];
Simulating model using a sample of the with a tremency of 1 minute
'and using optical input
watelletental. L
plt=0;
% sample schedule, for database has to meet, of lansute
t_insulin1_interp1 =
t_c_insulin1(1,1):1:Max_Exp_Dur; *t_s_accordent(ond,1);
%[min]
c_insulin1_interp1=interp1(t_c_insulin1(:,1),t_c_insulin1(:,2),...
    t_insulin1_interp1);
t_c_insulin1_interp1(:,1)=t_insulin1_interp1;
t_c_insulin1_interp1(:,2)=c_insulin1_interp1;
< simulating datal
[c_glucose_model_interp1,c_glucose_labeled_model_interp1,...
   c_interstitial_insulin_model_interp1,....
    c_interstitial_glucose_model_interp1,c_Ggut_model_interp1]=...
    gluc_sim_new(L0,t_c_insulin1_interp1, p_true,plt,...
    ·);
s sample schedules for data? with a treatency of 1 minute
t_insulin2_interp1 =
t_c_insulin2(1,1):1:Max_Exp_Dur; ' _____ (end, l);
121
c_insulin2_interp1=interp1(t_c_insulin2(:,1),t_c_insulin2(:,2),...
t_insulin2_interp1);
t_c_insulin2_interp1(:,1)=t_insulin2_interp1;
t_c_insulin2_interp1(:,2)=c_insulin2_interp1;
 % simulation using data,
    [c_glucose_model2,c_glucose_labeled_model2,...
       c_interstitial_insulin_model2,...
c_interstitial_glucose_model2,c_Ggut_model2,c_Ggut_labeled_model2]...
   * adding noise
datat_new=t_insulin1_interp1;
sdatal_new=awgn(c_ulucose_model_interpl,(1/0.06),'measured','linear');.
STAN RUIS
data1_new=c_glucose_model_interp1;
%data2_pew_awom(c_interstitial_rincose_ocdol2,(1/0.08),'measured',...
    'libear'); S+32 RUIS2 instructions
```

#### data2\_new=c\_interstitial\_glucose\_model2;

>data3\_new=awq+(c\_sturcsc\_tarched\_acuel\_nsterp1,(1.0.05),'measured',... % 'timear'); >=67\_E015% its constants data3\_new=c\_glucose\_labeled\_model\_interp1;

continuiali dish

#### i. t. 1

"daloulating the optimal samples include its insuit () and input 3 %Algorithms based on starting with a survestigal sample schedule where teach individual sample is optimiled by comparing it with its "Sneighbouring samples in the case is coal sample schedule.

```
%initialization de sonvertional samples chedule
opt_data1=data1;
opt_data3=data3;
opt_SS_data1=t_c_insulin1_interp1(:,1);
datac1_using=t_c_insulin1_interp1(:,2);
aa=1:1:(2*t_exp-l);
endvalue=1e6;
r=1;
for q=1*t_exp:t_exp:endvalue
    for z=[aa]
    qz=q+z karsplay or send a
    Establing with optimizing the stand to by appring P with its
    %neighbouring samples in the second could sample schedule.
    Essential single scherule ad its implitude
    datat1_using=opt_SS_data1;
    datat1_usinggg=opt_SS_data1' in () , a meric of
    data1_using=opt_data1;
    data3_using=opt_data3;
    datat1_using_old=datat1_using' de reavour menuter
    datat1_using(q/t_exp+1)=opt_SS_data1(q/t_exp)+z; craving bew sample
    .schedule
    datat1_using_new=datat1_using' = hearing to a source
    Sprocedure if the horizon is supposed to end it a sample which
    Surgitation that the last send to the suspinent of the search for
. . .
   % the optimal sample scheends : - erimes:
    if datat1_using(q/t_exp+1)>=Max_Exp_Durclit sample of experiment
       Sto ached
                                              : ')
       display('
       aa=1:1:(2-r)*t_exp-1 he hereth t the hereton must ...
       S at the stronger of the face out to
                                          3 K. (11)
```

```
r=r+1;
     if numel(aa)==0
                               . e ., net readed')
            display('
            Adefining the output comment of opt_sample_schedule_new
who ch
             is established to the or loop:
            opt_SS_data2=0
            opt_data2=0
            return
        end
    end
    data1_using_old=data1_using'; the old is represented.
    data1_using(q/t_exp+1)=data1_new(opt_SS_data1(q/t_exp)+z+1);
    savine has a finew sample:
    data3_using(q/t_exp+1)=data3_new(opt_SS_data1(q/t_exp)+z+1);
    <sup>9</sup> schedule
    data1_new_neww=data1_using'; on play removator
    data3_new_neww=data3_using'; capalay parameter
    datac1_using_old=datac1_using; discloy is notation
datac1_using(q/t_exp+1)=t_c_insulin1_interp1(opt_SS_data1(q/t_exp)+z+1,
2);
    Asaying new amplitudes of similar thempila:
    datat1_c_using=[datat1_using datac1_using];
   Rinterpl does not support the test of the as taken at a later time.
Than one of the tellemination (
    if (q/t_exp+2) <=length (datat1_using)</pre>
       if datat1_using(q/t_exp+1) == datat1_using(q/t_exp+2)
       pause
       ww=1;
       while (q/t_exp+1+ww) <=length (datat1_using) -1</pre>
        datat1_using(q/t_exp+1+ww) = datat1_using(q/t_exp+1+ww)+t_exp
        datac1_using(q/t_exp+1+ww) = datac1_using(q/t_exp+1+ww)+t_exp
        data1_using(q/t_exp+1+ww)=data1_using(q/t_exp+1+ww+1)
        data3_using(q/t_exp+1+ww)=data3_using(q/t_exp+1+ww+1)
       ww = ww + 1:
       end
        datat1_using=datat1_using(1:length(datat1_using)-1)
        datac1_using=datac1_using(1:length(datac1_using)-1)
        data1_using=data1_using(1:length(data1_using)-1)
        data3_using=data3_using(1:length(data3_using)-1)
        datat1_c_using=[datat1_using datac1_using];
        opt_SS_data1=datat1_using structure the current SS
        opt_data1=data1_using ? says the data of the current SS
        opt_data3=data3_using * saver the data of the current SS
   end
   end
```

```
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```

```
Seaturating parameters and of the new it corresponding tisher
    SHOULT
    [p_model,L0_model,c_glucose_model,c_glucose_labeled_model,...
    c_interstitial_insulin_model,...
    c_interstitial_glucose_model,c_Ggut_model,resnorm,residual,varp,...
stdp]=gluc_omm_mls_new(data1_using,data2,data3_using,L0,datat1_c_using,
. . .
    t_c_insulin2, a, b, c, sigma1, sigma2, sigma3, p_fix, p_init, t_exp, ...
    Nparameters,NoMean,plt);
    if z==aa(1)
             varp_old=varp; seeme he tars follow waters
             optimumDET=det(varp) i cela, se meser :
             opt_SS_data1=data1_using out a the area 35
             opt_datal=datal_using _____ area _ h __d d a ____ the further US
             opt_data3=data3_using
                                      aving the data of the current SS
             SAVED=1 with splay they
        elseif det(varp)<det(varp_old) stability the tisher matrix
                varp_old=varp; s is the invest determinant of the
Lisher
                 0.000
                SAVED=1
                                   to the the late
                optimumSample=qz at the real sector optimumDET=det(varp) and the monitor
                opt_SS_data1=datat1_using saving the current SS
                opt_data1=data1_using ' saving the data of the current
351
                                       even the diffict the current
                opt_data3=data3_using
13
        else
            det (varp) seudellas, et americanes
            NOTSAVED=1 display of toutor
        end
         end
end
end
1 f 0
scalculating the optimal sample of the high for rapht 2
Silutialization of conventional sample schedule.
sAlgorithms based or startand with a traventional sample schedule where
Seach individual sample is a time of comparing of with its
speighbouring complex in the cover of sample schedule.
siduralization of some rest and standard
```

```
opt_data2=data2;
```

```
opt_SS_data2=t_c_insulin2(:,1);
```

```
aa=t_exp:t_exp:5*t_exp;
endvalue≈1e6;
r=1;
for g=1*t_exp:t_exp:endvalue
   for z=[aa]
   qz=q+z %diselsy or monitor
   scharting with optimal is the fis sumply by comparing it with it.
   sheighbon in ample i be sveri yet ample acherube.
   Saving optimal sample schedule as it samplitude.
   datat2_using=opt_SS_data2;
   data2_using=opt_data2;
      calculating the mean of the interpolated datas at the data2
-amples
   SumOfData=0;
   if NoMean==0
   for pp=1:1:t_exp
   SumOfData=SumOfData+data2_new(opt_SS_data2(q/t_exp)+z+1-(pp));
   Ssimming data of new same b
   And
   Meandata=1/t_exp*SumOfData; Mean of severamele
   data2_using(q/t_exp+1)=Meandata; saturd data of rew sample
   else
   . . .
   data of pew ample
   end
   datat2_using_old=datat2_using' realises a meritor
   datat2_using (q/t_exp+1) = opt_SS_data2 (q/t_exp) +z; set and new sample
    acheoula
   data2_new_neww=data2_using' in the second second
   Recordered of the begins of such as a set of sense at a sample which
1 . . .
   sablated that the last cample in the second of the second for
    the optimal sample schedule is the bed
   if datat2_using(g/t_exp+1)>=Max_Exp_Dur class sample of experiment
      Steached
                                          · · · · )
       display('
      aa=t_exp:t_exp:(5-r)*t_exp; the performance the housen must ...
      3 at least end at the last semile taken
      r=r+1;
      if numel(aa)==0
           display('
                                        a. I i he ')
           %detrning the cutput of uncent cort opt_sample_schedule_new
which
           sare not calculated in this ta loop:
           opt_SS_data1=0
           opt_SS_data3=0
           opt_data1=0
```

# D.7 White Noise test

#### wnoisetest\_new.m

```
function [datacov,mean_data1,mean_data2,mean_data3]=...
    wnoisetest_new(data1, data2, data3, residual)
0,00
%Oral Minimal Model of glucose kipetics TOOL
%This function analyzes the properties of data used en residuals leit
3during the estimation procedure
0.00
2
%History
307 july 06, Salvador Almagre Frutes, TU/e
0, 000
°Function:
%[datacov_data1,mean_data1,mean_data2]=wneisetest(data1,data2,residual)
%Inputs function:
%datal: data of interstitual insulin [pM]
%data2: data of interstitial glucose [mM]
idatai: data of labeled plasma clucese [mM].
Stepidual: the value of the residual of the estimation algorithm.
2
%Outputs function:
%datacov_datal: crosscorrelation of the data.
%mean_datal; calculating the mean of the data. Must be zero.
%mean_data2: calculating the mean of the data. Must be zero.
%mean_data3: calculating the mean of the data. Must be zero.
% TEST: calculating the mean. Must be vere.
mean_data1=mean(data1);
mean_data2=mean(data2);
mean_data3=mean(data3);
\% TEST: Statistical properties: gaussime distribution with zero mean
pp=-1:0.01:1;
yy=hist(residual,pp);
figure;
plot(pp,yy);title('
                                   );
                   `);
xlabel('
ylabel('
                   ');
% normalizing residuals
residual=residual./max(residual);
% TEST: crosscorrelation. Must be the form of a pulse at zero
timeshift:
for m=0:(length(residual)-round(0.4*length(residual)))
    sum=0;
```

```
for n=1:(round(0.4*length(residual)))
    part=residual(n)*residual(n+m);
    sum=sum+part;
    end
datacov(m+1)=(1/(round(0.4*length(residual)))).*sum;
end
figure
plot((0:(length(residual)-round(0.4*length(residual)))),datacov);
title(' = = = = ');
xlabel(' = = = ');
ylabel(' = = = ');
```

## D.8 Worst case bias S<sub>1</sub>-cap

# D.8.1

## worstcase\_bias\_steil.m

y c o %Oral Minimal Model of glucose kinetics. 2 %This script estimates the identiafiable parameters in Steil's %Model of insulin kinetics using Maximum least %Squares Estimation õ 0,29 20 %Historγ \$28 July-06, Salvador Almagro Frutos, TU/o 9,29 Binitializes all parameters of Sterl's Model of insulin kintics hand all time samples of plasma insulin. clear gluc\_param\_new; Nparameters=length(p\_true); NoMean=0;SCalculates the mean over t\_exp minutes of interstitial alucose %samples if NoMean=0 diary 1 1 % Sets the names of the mat files containing the estimated parameters: % make plots if plt=1 plt≈1; NoMean==0 Interpolating data2. t\_insulin2\_interp1 = t\_c\_insulin2(1,1):1:t\_c\_insulin2(end,1);%[min] c\_insulin2\_interp1=interp1(t\_c\_insulin2(:,1),t\_c\_insulin2(:,2),... t\_insulin2\_interp1); t\_c\_insulin2\_interp1(:,1)=t\_insulin2\_interp1; t\_c\_insulin2\_interp1(:,2)=c\_insulin2\_interp1; % simulation using data2 [c\_glucose\_model2\_interp1, c\_glucose\_labeled\_model2\_interp1, ... c\_interstitial\_insulin\_model2\_interp1, ... c\_interstitial\_glucose\_model2\_interp1,... c\_Ggut\_model2\_interp1,c\_Ggut\_model2\_labeled\_interp1].... 

```
% calculating the mean of the interpolated data? at the data?
samples
    GIb=L0(4);
    c_interstitial_glucose_mean(1,1)=GIb;
    i=2;
    for r=t_c_insulin2(2:end,1)';
    c_interstitial_glucose_mean(i,1) = ...
    sum(c_interstitial_glucose_model2_interp1(r+1-t_exp:r))/t_exp;
    i=i+1;
    ena
    else
    % simulation using data2
    [c_glucose_model2, c_glucose_labeled_model2, ...
        c_interstitial_insulin_model2,...
c_interstitial_glucose_model2,c_Ggut_model2,c_Ggut_labeled_model2]...
    =gluc_sim_new (L0,t_c_insulin2,p_true,plt,''decomposition of );
    end
% adding noise
if NoMean==0
data2_new=c_interstitial_insulin_model2_interp1;
%data2_new=data2_new+0.10 * randn(l+nuth(data2_new),1);%+8% RUIS
4 50
data2_new=c_interstitial_insulin_model2;
%data2_new=data2_new+0.10 * rands(length(data2_new),1); *+8% RUIS%
ena
% make plots if plt-1
plt=1;
steil_param;
Nparameters=length(p_true);
% End initialisation
% 2-dimensional parameter space
%p: = [Ib SIcap p2];
%1 known parameters:
p_fix = [Ib p2];
%1 unknown model parameters:
p_init = [SIcap]
p_init_copy=p_init;
estimation(1,:)=[-0.5 - 0.4 - 0.3 - 0.2 - 0.1 0.1 0.2 0.3 0.4 0.5];
estimation1(1,:)=estimation(1,:);
for kk=1:length(p_init)
    11=0;
    for zz=estimation(1,:)
    p_init=p_init_copy;
     11=11+1;
     estimator=find(p_init(kk)==p_true);
```

```
p_init(kk)=p_init(kk)+ zz*p_init(kk)
     %estimation of parameters
[p_model,L0_model,c_interstitial_insulin_model,resnorm,residual,varp,...
stdp]=gluc_omm_mls_steil(data2_new,L0,t_c_insulin2_interp1,a,sigma1,...
   p_fix,p_init,Nparameters,plt);
   p_Model=p_model(estimator) &display on monitor
   estimation(kk+1,ll)=p_model(estimator);
   estimation1(kk+1,ll)=p_model(estimator1);
   end
end
        TARGET=fopen(filename, ' ');
        fprintf(TARGET,...
           • • .
                                   . . 1 * 1 . L.* 10.1** *11.10f
        1.1
                             ',estimation1');
                1
       fclose(TARGET);
```

estimator1=find(p\_init(1)==p\_true);

p\_True=p\_true(estimator) sdisplay on monitor

#### steil\_param.m

مان مان % Steil's Model of insulin kinetics. Š %This script initializes all parameters of Steil's Model of insulin %kinetics and the time samples of plasma insulin. olo ol × %History %28-July-06, Salvador Almagro Frutos, 10/e %29-Jan-04, Natal van Riel, TU/e %17-Jun-03, Natal van Riel, TU/e of of %model parameters: %Values from: 2?????? Ib = Ib; %[pM] baseline insulin concentration in plasma **SIcap = SIcap;** %[1/min] transendothelial insulin transport parameter p2 = p2; [1/an] total fractional insulin clearance from the %interstitial fluid p\_true = [Ib, SIcap, p2];% 2 dimensional parameter space %Fixed initial conditions clear LJ; LO(1) = 0; % [pM] insulin concentration in interstitium Eleput insulin concentration in plasma [pM]; assumed to be known at each ssimulation time sample from linear interpolation of its measured samples %Time and amplitude samples of plasma insulin: %l-> blood glucose data t\_insulin1 = t\_c\_insulin1(:,1); c\_insulin1 = t\_c\_insulin1(:,2); t\_c\_insulin1(:,1) = t\_insulin1'; t\_c\_insulin1(:,2)=c\_insulin1'; % squared 2-norm of the error-function residuals a=1; %output weighting factor of plasma glucose concentration (lst output) sigmal=1; %variance of plasma glucose concentration(lst output)

```
%Plots
% make plots if plt=1
if plt==1;
%Plots
```

### steil\_sim.m

```
function [c_interstitial_insulin]=steil_sim
(L0,t_c_insulin,p,plt,TITLE)
9. 9. 9
Steil's Model of insulin kinetics.
%This function makes a simulation of the Sterf's Model of insulin
%kinetics using the parameters defined is sterl_param.m.
مان
مان
مان
%Function:
%[c_interstitial_insulin]=sterl_sim (L0,t_c_insulin1,t_c_insulin2,p,...
%plt,TITLE)
%Inputs function:
%L0: Fixed initial conditions,L0(1) - [pM] insulin concentration in
%interstitium.
t_c_insulm: [pM] [min] Time and amplitude samples matrix of plasma
%insulin 1 > blood glucose data 2 \rightarrowinterstitual glucose data
%p: = [SIcap, p2];% 2 dimensional parameter space
%plt: plots are made in this function if plt=1.
<code>%TITLE:</code> set the title-name if plots are made in this function (plt=1).
5
%Outputs function:
%c_interstitial_insulin: [pM] interstitial insulin concentration.
of conc
ç
%History
%28-July-06, Salvador Almagro Frutos, TU/e
%29-Jan-04, Natal van Riel, TU/e
%17 Jun 03, Natal van Riel, TU/e
دیر
تېر
%model parameters:
%Slcap [1/min] transendotheliai insulin transport parameter
%p2 (1/min) total fractional insulin plearance from the
%interstitial fluid
%Differential equations solver: ode15s
ode_options =odeset('F+ I '', le-6, ' + F', le-12);% parameters 2-end
%%ode_options =odeset('RelTol', le=9, 'AbsTol', le 16);
[t,L] = ode45(@steil_ode,t_c_insulin(:,1),L0,ode_options,
t_c_insulin,p);
*Out put
c_interstitial_insulin=L(:,1); %[pM] interstitial insulin concentration
% make plots i! pit-1
```

```
it plt==1
3Plots
figure (333);
Ib=p(1);
subplot(221);plot(t_c_insulin(:,1),t_c_insulin(:,2),' ',
' an wat h',2);
hold ;
%Baseline level:
plot( [t_c_insulin(1) t_c_insulin(end)], [Ib Ib], '
Stigare
subplot(222); plot(t_c_insulin(:,1), c_interstitial_insulin,
' a + (+ (b',2)
            '); ylabel...
xlabel('at the state
                                · · · · · )
 ( '
Title(TITLE);
```

end

# steil\_ode.m

```
function dLout = steil_ode(t,Lin,t_c_insulin,p)
999
%STELL_ODE ODE's of STELL'S Model of insulin kinetics.
2
%This function sets STELL_ODE ODE's on STELL'S Model of insulin
%kinetics using the parameters defined in sterl_param.m.
510
ŝ
%Function:
%dLout - sterl_ode(t,Lin,t_c_insulin1,p).
%inputs tunction:
%Lin: state values at previous time sample L(k, 1) = [X(k, 1)]
%t_c_insulin: c_insulin(k) (input sideal at sample k) OK matrix
%t_c_insulın
%p: = [Slcap, µ2];% 2 dimensional parameter space
ý.
%Outputs function:
%dLout: new state derivatives dL(k) [dX(k);j
CiC
of of of
¥.
SHistory
%28-July-06, Salvador Almagro Frutos, TU/e
%29-Jan-04, Natal van Riel, TU/e
%17-Jun-03, Natal van Riel, TU/e
UN CIO
% defining the row were each state in statevecor Lin is present
idX = 1; %interstitial insulin concentration [pM]
% model parameters:
Ib = p(1); %[pM] baseline insulin concentration in plasma
SIcap = p(2); %[1/min] transendothelial insulin transport parameter
p2 = p(3);  [1/min] total fractional insulin clearance from the
finterstitial fluid
%calculate c_insulin by interpolation of t_c_insulin
u = interp1(t_c_insulin(:,1),t_c_insulin(:,2), t);
%ode's
dX = SIcap/1e3*(u-Ib)^2-p2*Lin(idX);
```

dLout = [dX];

### gluc\_omm\_mls\_steil.m

```
function [p_model,L0_model,c_interstitial_insulin_model,...
resnorm, residual, varp, stdp]=gluc_omm_mls_steil...
(data1,L0,t_c_insulin1,a,sigma1,p_fix,...
p_init,Nparameters,plt)
0000
%STEIL MM MIS Maximum Least Squares Estimation of Steil's Model of
%insulin kinetics
õ
000
<sup>§</sup>This function estimates the parameters using Maximum Least Squares
%Fstimation of Steil's Model of insuler kinetics using the
Sparameters defined in stell param.m.
600
%Function:
%[p_model,10_modelc_interstitial_insulin_model,...
%resnorm,varp,stdp]=qluc_omm_mls_steil...
%(data1,L0,t_c_insulint,p_tix,p_init,Nparameters,plt)
<u>ç</u>
">Inputs function:
Sdatal: data of interstitual insulin [pM]
"L0: Fixed initial conditions,L0(1)= [pM] insulin concentration in
Einterstitium.
%t_c_insulin: [pM] [min] Time and amplitude samples matrix of plasma
%insulin 1-> blood glucose data
%p: = [Slcap, p2];% 2-dimensional parameter space
%p_fix: parameter space which is fixed.
'sp_init: initialisation of parameter space which will be estimated.
"sampleinterval [min] of data 2.
*Nparamters: number of dimensional parameter space p.
%plt: plots are made in this function if pl*-1.
%Outputs function:
%p_model: [S[cap,p2];% 2 dimensional parameter space
%c_interstitial_insulin_model: [pM] interstitial insulin concentration.
%resporm: squared 2-norm of the error function residual.
%residual: the value of the residual of the estimation algorithm.
%varp: variance
%stdp: The standard deviation is the square root of the variance.
0,00
0
SHistory
528-July 06, Salvador Almagro Frutos, PY/e
%29-Jan-04, Natal van Riel, TU/e
%17 Jun 03, Natal van Riel, TU/e
0.00
% lower (1b) and upper (ub) bounds, so that the solution of the
estimated.
\% parameters is always in the range 1b \sim parameter <- ub.
```

```
lb = [];%0*ones(size(p_init));%[0 0 0 0];
ub = [];%ones(size(p_init));%[];
% provides the function specific details for the options values for
% lsqnonlin
options =optimset(' : ; !...,',' ',' ', 'e-10,... %'iter'
default:le-4
'] [//,1e-20,... %default: le 4
'Maxhu Eval', [800],... %800 default:
'Maxhu Eval', [800],... %800 default:
'Marhu Eval', ',' ',... %default: on
'Marhu Scathi', ',' (that 1 ');
%'Jacobian', 'on');
%LSQNONLIN: objective function should return the model error
%p_estimated: estimated parameters
%resnorm: squared 2-norm of the residual
tresidual: returns the vector of the residuals
\beta_{\text{Jacobian}}: returns the Jacobian of tun at the solution x.
[p_estimated, resnorm, residual, exitflag, output, lambda, jacobian]...
   =lsqnonlin(@obj_fn_steil,p_init,lb,ub,options,p_fix,L0,datal,...
    t_c_insulin1, a, sigma1, Nparameters);
if exitflag==0
STOP=1;
end
%Accuracy:
%lsqnonlin returns the jacobran as a sparse matrix
varp = resnorm*inv(jacobian'*jacobian)/length(t_c_insulin1(:,1));%
variance
stdp = sqrt (diag(varp)); "The standard deviation is the square root
%of the variance
p_model = [p_fix(1) p_estimated(1) p_fix(2)];
L0_model = [L0(1)];
% make plots with new 2 dimensional parameter space if plt=1
[c_interstitial_insulin_model] = steil_sim (L0,...
    t_c_insulin1,p_model,plt,
                                                ');
```

# obj\_fn\_Steil.m

function error = obj\_fn\_Steil(p\_var, p\_fix, L0, data1,t\_c\_insulin1... ,a,sigma1,Nparameters) 010 010 010 SObjective function Maximum Least Squares Estimation of Stell's %Model of insulis kinetics. 5 80,00 "This function is used as objective function for LSQNONLIN. 510 210 \*Function: %e obj\_fn\_steil(p\_var, p\_fix, L0, datal, \_\_\_\_insulin),... 02 Npalameters) ò #Inputs function: %p\_var: parameter space which will be estimated %p\_fix: parameter space which is fixed. %L0: Fixed initial conditions, L0(1)- [pM] insulin concentration in ... %interstitium. %data1: data of interstitual insulum [pM]. %t\_2\_insulin: [pM] [min] Time and amplitude samples matrix of plasma Sinsulin. sa: output weighting lactor of 1st output. Ssigmal: variance of 1st output. %Nparamters: number of dimensional parameter space p. estimator: defines the parameter which is used in the model to model %adjustment identification. 1 %Outputs function: Serror: minimization variable of MLS (molel error) 19.6 Q. %History %28 July-06, Salvador Almagre Frutes, IU/e %29 Jan-04, Natal van Riel, TU/e S17 Jun C3, Natal van Riel, TU e

### p\_estimating=[p\_fix(1) p\_var(1) p\_fix(2)];

%estimating 2-dimensional parameter space %p\_true = [Ib,SIcap,p2];

%Simulating Stell model of insulin kinetics using the estimating %2-dimensional parameter space and the interpolated insulin concentration %samples for datal.

```
plt=0; % make NO plots
[c_interstitial_insulin_estimating]=...
```