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**A Generalized Memoryless Property**

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O. Kella, A. Löpker  
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# A Generalized Memoryless Property

Offer Kella<sup>\*†</sup>      Andreas Löpker<sup>‡</sup>

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## Abstract

We consider a generalized memoryless property which relates to Cantor's second functional equation, study its properties and demonstrate various examples.

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**AMS Subject Classification:** Primary: 60K30 Secondary: 90B18, 60J25, 60J35, 60G44, 68M20.

## 1 Introduction and setup

Let  $\mathbb{R}$  denote the set of real numbers and  $\mathcal{B}$  be it's Borel sets. Consider a Markov kernel  $P(x, A)$  where for each  $x \in \mathbb{R}$ ,  $P(x, \cdot)$  is a probability measure on  $(\mathbb{R}, \mathcal{B})$ . We say that the generalized memoryless property is satisfied if the following is satisfied for each nonnegative  $x, y$  and real  $z$ :

$$P(z, (x + y, \infty)) = P(z, (x, \infty))P(z + x, (y, \infty)) \quad (1)$$

We note that if  $P(x, \cdot) = P(\cdot)$ , that is, the Markov kernel is independent of the originating state, this is precisely the memoryless property of which the only solution is of the form  $P((x, \infty)) = \rho^x$  for some  $0 \leq \rho \leq 1$ . That is, a random variable that has such a distribution is either almost surely (a.s.) zero, or a.s. infinite or has an exponential distribution.

To motivate the property given by (1) assume that in addition  $P(\cdot, A)$  is a Borel function for each  $A \in \mathcal{B}$  and let  $T_a$  (age) and  $T_r$  (remaining life) be a pair of random variables, where  $T_a$  has an arbitrary distribution and  $\mathbb{P}(T_r \in A | T_a) = P(T_a, A)$ . That is, one interprets  $\mathbb{P}(T_r \in A | T_a = z) = P(z, A)$ . Then (1) becomes

$$\mathbb{P}(T_r > x + y | T_a = z) = \mathbb{P}(T_r > x | T_a = z) \mathbb{P}(T_r > y | T_a = z + x) . \quad (2)$$

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<sup>\*</sup>Department of Statistics; The Hebrew University of Jerusalem; Mount Scopus, Jerusalem 91905; Israel (Offer.Kella@huji.ac.il).

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<sup>‡</sup>Department of Economics and Social Sciences; Helmut Schmidt University Hamburg, 22043 Hamburg; Germany (lopker@hsu-hh.de.de).

In other words, in order for the a component having age  $z$  to function at time  $x + y$ , it first has to function at time  $x$ . Then, independent of everything else, its age is modified to  $z + x$  and, given that its age is now  $z + x$ , it has to function at time  $y$ . This model relates to *Model II* of [7]. In fact, given the condition that  $P(0, (x, \infty)) > 0$  for each  $x \geq 0$  this results in exactly the same model, even though we write it in more primitive terms. To see this, denote  $1 - F(x) = P(0, (x, \infty))$  and observe that with  $z = 0$ , (2) becomes

$$1 - F(x + y) = (1 - F(x))\mathbb{P}(T_r > y | T_a = x), \quad (3)$$

so that indeed

$$\mathbb{P}(T_r > y | T_a = x) = \frac{1 - F(x + y)}{1 - F(x)}, \quad (4)$$

which is equation (1) of [7] with the roles of  $x$  and  $y$  reversed. However, if we only assume that  $P(z, (x, \infty)) > 0$  for all  $0 < z \leq x$  then we will see that it is possible that  $\mathbb{P}(T_r > y | T_a = x) = s(x + y)/s(x)$  for some nonincreasing function  $s$  for which  $s(x) \rightarrow \infty$  as  $x \downarrow 0$ . We will also see that other possibilities may occur as well. We also mention that for the case where  $P(z, (x, \infty)) > 0$  a more general model was considered in [9]. However, in this paper the authors suffice in pointing out some examples of this property but do not characterize the general form. In this generality, a characterization might not be possible.

To continue, for real  $x$  and  $y \geq x$  we denote  $\mu(x, y) = P(x, (y - x, \infty))$  and observe that (1) becomes

$$\mu(z, y) = \mu(z, x)\mu(x, y) \quad (5)$$

for each  $z \leq x \leq y$ . If this was valid for all  $x, y, z$  then (5) is called *Cantor's second functional equation*. We note that in the latter case, if  $\mu(z, y) \neq 0$  for some  $z, y$  then  $\mu(z, x) \neq 0$  and  $\mu(x, y) \neq 0$  for all  $x$ . Since  $\mu(x, y) = \mu(x, u)\mu(u, y)$  then  $\mu(x, u) \neq 0$  for all  $x, u$  and in particular if we denote  $s(x) = \mu(z, x)$  then for any  $x, y$  we have that  $\mu(x, y) = s(y)/s(x)$ . Thus, as is well known (e.g., see [1]), the only solutions of Cantor's second functional equation are either  $\mu(x, y) = 0$  for all  $x, y$  or  $\mu(x, y) = s(y)/s(x)$  for some function  $s(\cdot)$  that never vanishes. In the second case it is clear that for all  $x$ ,  $\mu(x, x) = 1$  and that for any nonvanishing function  $s(\cdot)$ ,  $s(y)/s(x)$  obeys Cantor's second functional equation.

When we assume that (5) is satisfied only if  $z \leq x \leq y$ , the solution requires a bit more care.

## 2 Main observations

In order to consider the most general setup, let us assume until further notice that

$$\mu : \{(x, y) | x \leq y\} \rightarrow \mathbb{R} \quad (6)$$

(note:  $\mathbb{R}$  rather than just  $[0, 1]$ ) and that (5) is satisfied for  $z \leq x \leq y$ . Denote

$$b(z) = \begin{cases} z & \text{if } \mu(z, x) = 0 \ \forall x > z \\ \sup\{x | \mu(z, x) \neq 0, x > z\} & \text{otherwise} \end{cases} \quad (7)$$

and

$$a(z) = \begin{cases} z & \text{if } \mu(x, z) = 0 \forall x < z \\ \inf\{x | \mu(x, z) \neq 0, x < z\} & \text{otherwise.} \end{cases} \quad (8)$$

Now, when  $a(z) < b(z)$  we denote

$$I(z) = \begin{cases} (a(z), b(z)) & \text{if } \mu(a(z), z) = 0 = \mu(z, b(z)) \\ (a(z), b(z)] & \text{if } \mu(a(z), z) = 0 \neq \mu(z, b(z)) \\ [a(z), b(z)) & \text{if } \mu(a(z), z) \neq 0 = \mu(z, b(z)) \\ [a(z), b(z)] & \text{if } \mu(a(z), z) \neq 0 \neq \mu(z, b(z)) \end{cases} \quad (9)$$

and when  $a(z) = b(z) = z$  we let  $I(z) = \{z\}$ , noting that from  $\mu(z, z) = \mu(z, z)\mu(z, z)$  necessarily  $\mu(z, z) \in \{0, 1\}$ . Moreover, since

$$\mu(u, v) = \mu(u, v)\mu(v, v) = \mu(u, u)\mu(u, v) \quad (10)$$

for  $u \leq v$ , it is easy to check that if  $I(z)$  is not a singleton, then necessarily  $\mu(z, z) = 1$ .

Consider now the following.

**Lemma 1.** *Let  $\mu : \{(x, y) | x \leq y\} \rightarrow \mathbb{R}$  satisfy (5) for all  $z \leq x \leq y$ . Then for  $x < y$ ,  $\mu(x, y) \neq 0$  if and only if  $I(x) = I(y)$  and for every  $u, v \in I(x)$  with  $u \leq v$  we have that  $\mu(u, v) \neq 0$ .*

*Proof.* For  $v \geq y$  we have that  $\mu(x, v) = \mu(x, y)\mu(y, v)$ , so that  $\mu(x, v) \neq 0$  if and only if  $\mu(y, v) \neq 0$ . For  $u \leq x$  we have by similar reasoning that  $\mu(u, x) \neq 0$  if and only if  $\mu(u, y) \neq 0$ . For  $w \in (x, y)$  we have that  $\mu(x, y) = \mu(x, w)\mu(w, y)$  and thus  $\mu(x, w) \neq 0$  and  $\mu(w, y) \neq 0$ . This implies that  $I(x) = I(y)$ . Now, for every  $u \in I(x)$  we have that  $I(u) = I(x)$  and thus for any  $v \geq u$  with  $v \in I(x)$  it follows that  $v \in I(u)$  which implies that  $\mu(u, v) \neq 0$ .  $\square$

**Theorem 1.** *Under the conditions of Lemma 1 there exists a family  $\{I_\theta | \theta \in \Theta\}$  of necessarily at most countably many disjoint intervals (open, half open or closed) and possibly uncountably many singletons with  $\cup_{\theta \in \Theta} I_\theta = \mathbb{R}$  such that for each  $\theta$  for which  $I_\theta$  is not a singleton there exists a function  $s_\theta : I_\theta \rightarrow \mathbb{R}$  which is nonvanishing such that for every  $x, y \in I_\theta$  with  $x \leq y$  we have that  $\mu(x, y) = s_\theta(y)/s_\theta(x)$ .*

*Proof.* For an arbitrary  $z \in I_\theta$  define

$$s_\theta(x) = \begin{cases} \mu(z, x) & \text{if } z \leq x \in I_\theta \\ 1/\mu(x, z) & \text{if } z \geq x \in I_\theta. \end{cases} \quad (11)$$

Then, if  $z \leq x \leq y$  and  $x, y \in I_\theta$  then

$$s_\theta(y) = \mu(z, y) = \mu(z, x)\mu(x, y) = s_\theta(x)\mu(x, y). \quad (12)$$

If  $x \leq y \leq z$  and  $x, y \in I_\theta$  then

$$\frac{1}{s_\theta(x)} = \mu(x, z) = \mu(x, y)\mu(y, z) = \mu(x, y)\frac{1}{s_\theta(y)}. \quad (13)$$

Finally, when  $x \leq z \leq y$  with  $x, y \in I_\theta$  then

$$\mu(x, y) = \mu(x, z)\mu(z, y) = \frac{1}{s_\theta(x)}s_\theta(y). \quad (14)$$

□

In particular we observe that if  $\mu(x, y) \neq 0$  for all  $x \leq y$  then  $I(x) = \mathbb{R}$  and for some nonvanishing function  $s : \mathbb{R} \rightarrow \mathbb{R}$  we have that  $\mu(x, y) = s(y)/s(x)$  for all  $x \leq y$ . It seems as if this is the same solution of Cantor's second functional equation until we recall that  $\mu(x, y)$  is undefined for  $x > y$ .

Returning to the generalized memoryless property of (1) we recall that  $\mu(x, y) = P(x, (y - x, \infty))$  and thus, for  $y \geq 0$ ,  $P(x, y) = \mu(x, x + y)$ . Hence we conclude the following.

**Theorem 2.** *Assume that (1) is satisfied. Then there exists a family  $\{I_\theta \mid \theta \in \Theta\}$  of disjoint intervals and singletons with  $\cup_{\theta \in \Theta} I_\theta = \mathbb{R}$  such that for each  $\theta$  for which  $I_\theta$  is not a singleton there exists a nonincreasing right continuous strictly positive function  $s_\theta : I_\theta \rightarrow \mathbb{R}$  such that for every  $x \in I_\theta$  and  $y \geq 0$  with  $x + y \in I_\theta$  we have that*

$$P(x, (y, \infty)) = \frac{s_\theta(x + y)}{s_\theta(x)}.$$

**Remark 1.** We note that the same result holds, only with  $s_\theta$  being left continuous, for  $P(x, [y, \infty))$  (left closed interval) if we replace (1) by the left closed version

$$P(z, [x + y, \infty)) = P(z, [x, \infty))P(z + x, [y, \infty)) \quad (15)$$

which, by taking  $y \downarrow u$ , where  $u \geq x$ , is also equivalent to

$$P(z, (x + u, \infty)) = P(z, [x, \infty))P(z + x, (u, \infty)) \quad (16)$$

whenever  $x, y \geq 0$ .

**Remark 2.** For the case  $\mu(x, y) = P(x, (y - x, \infty))$  clearly  $\mu(x, y) \in [0, 1]$  for all  $x \leq y$ . Therefore, from  $\mu(z, y) = \mu(z, x)\mu(x, y) \leq \mu(x, y)$  for  $z \leq x \leq y$  it follows that  $\mu(\cdot, y)$  is a nondecreasing function on  $(-\infty, y]$ . Similarly  $\mu(x, \cdot)$  is nonincreasing on  $[x, \infty)$ .

**Remark 3.** It is easy to check that these results remain unchanged if the domain of  $\mu$  is  $0 \leq x \leq y$  rather than  $x \leq y$  or  $x \leq y \leq \infty$  or any other combination like this. Basically, whenever we have (5) for  $z \leq x \leq y$  these results apply.

**Remark 4.** When  $P(x, (y, \infty)) > 0$  for every  $x$  and every  $y \geq 0$ , then there is some necessarily nonincreasing and right continuous function  $s : \mathbb{R} \rightarrow (0, \infty)$  for which  $\mu(x, (y, \infty)) = s(x + y)/s(x)$  for all  $x$  and  $y \geq 0$ . Moreover the function  $s$  is uniquely determined up to a constant multiple. We note that unlike in [7], it is possible that for some  $\theta$ ,  $I_\theta = (0, \infty)$  and  $s_\theta(x) \rightarrow \infty$  as  $x \downarrow 0$ . If  $[0, \infty) \subset I_\theta$  for some  $\theta$  then one has the model considered in [7].

**Remark 5.** It is also evident that the same structure holds when one reverses (1) to

$$P(z, (-\infty, x + y)) = P(z, (-\infty, x))P(z + x, (-\infty, y)) \quad (17)$$

or to

$$P(z, (-\infty, x + y]) = P(z, (-\infty, x])P(z + x, (-\infty, y]) \quad (18)$$

for all  $x, y \leq 0$ . In particular when  $P(z, (-\infty, 0)) = 0$  and  $-z \leq x, y \leq 0$ . We will use this in the next section when modeling a certain growth collapse (additive increase multiplicative decrease) process with state dependent decrease ratios.

Consider now a possibly infinite interval  $I_\theta$  which is not a singleton and its corresponding positive valued function  $s_\theta$ . Then for each  $z \in I_\theta$  and  $x \geq 0$  such that  $z + x \in I_\theta$  we have that  $P(z, (x, \infty)) = s_\theta(z + x)/s_\theta(z)$ . If  $x + z \notin I_\theta$  then  $P(z, (x, \infty)) = 0$  so that we may define  $s_\theta(y) = 0$  for any  $y \notin I_\theta$  which is on the right of  $I_\theta$  (if any) where necessarily  $s_\theta$  must be right continuous at  $z^*(\theta) = \sup\{z \mid z \in I_\theta\}$ . Now, for  $z \in I_\theta$  and  $0 < u < 1$  denote

$$\begin{aligned} t_\theta^z(u) &= \inf \left\{ x \mid 1 - \frac{s_\theta(z + x)}{s_\theta(z)} \geq 1 - u, x \geq 0 \right\} \\ &= -z + \inf \{ x \mid s_\theta(x) \leq s_\theta(z)u, x \geq z \} \\ &= -z + \inf \{ x \mid s_\theta(x) \leq s_\theta(z)u \} \end{aligned} \quad (19)$$

Where the last equality follows since  $s_\theta(x) < s_\theta(z)u$  for every  $x \leq z$ . It is standard that if  $U \sim \text{Uniform}(0, 1)$  then  $t_\theta^z(1 - U)$  and thus  $t_\theta^z(U)$  have the distribution  $P(z, \cdot)$ . Thus, if we denote  $t_\theta(v) = \inf\{x \mid s_\theta(x) \leq v\}$  for  $v > \inf\{z \mid z \in I_\theta\}$  then for every  $z \in I_\theta$  we have that  $t_\theta(s_\theta(z)U) - z$  has the distribution  $P(z, \cdot)$ . Recalling  $T_a$  and  $T_r$  from Section 1, this implies the following.

**Theorem 3.** *Assuming that  $T_a$  and  $U \sim \text{Uniform}(0, 1)$  are independent, then  $(T_a, T_r)\mathbb{1}_{\{T_a \in I_\theta\}}$  and  $(T_a, t_\theta(s_\theta(T_a)U) - T_a)\mathbb{1}_{\{T_a \in I_\theta\}}$  are identically distributed.*

Clearly, when  $\Theta$  is countable then the immediate conclusion is that

$$(T_a, T_r) \sim \left( T_a, \sum_{\theta} (t_\theta(s_\theta(T_a)U) - T_a)\mathbb{1}_{\{T_a \in I_\theta\}} \right). \quad (20)$$

It is interesting to check when for a given  $\theta$  for which  $I_\theta$  is not a singleton the value of  $t_\theta(s_\theta(z)u) - z$  is independent of  $z$ . That is, it is only a function of  $u$ . The answer is not surprising.

**Theorem 4.** *When  $I_\theta$  is not a singleton then  $t_\theta(s_\theta(z)u) - z$  is independent of  $z \in I_\theta$  if and only if for some  $0 \leq \lambda_\theta < \infty$  and  $0 < c_\theta < \infty$ ,  $s_\theta(z) = c_\theta e^{-\lambda_\theta z}$  for  $z \in I_\theta$ .*

*Proof.* Let  $f(u) = t_\theta(s_\theta(z)u) - z$  (independent of  $z$ ) for every  $z \in I_\theta$  and  $0 < u < 1$ . Note that since the right side is left continuous in  $u$ , then so is  $f$  (as a function defined on

$(0, 1)$ ). In particular  $f$  is Borel. Denoting  $X = f(U)$  we have that for every  $z \in I_\theta$  and every  $x, y \geq 0$ , with  $z + x + y \in I_\theta$ ,

$$\begin{aligned} \mathbb{P}(X > x + y) &= \frac{s_\theta(z + x + y)}{s_\theta(z)} \\ &= \frac{s_\theta(z + x)}{s_\theta(z)} \cdot \frac{s_\theta(z + x + y)}{s_\theta(z + x)} \\ &= \mathbb{P}(X > x)\mathbb{P}(X > y) \end{aligned}$$

The equation  $g(x + y) = g(x)g(y)$  for  $x, y \geq 0$  under minor regularity conditions on  $g$  implies that  $g$  is either identically zero, identically one or exponential. Monotonicity, right or left continuity or even Lebesgue measurability are sufficient conditions. The standard proof can be easily modified to the case where  $g$  is defined on and the equation is valid only when  $x, y, x + y$  are in  $[0, a)$  or  $[0, a]$  for some  $0 < a < \infty$ , resulting in  $g$  being identically zero, identically one or exponential on  $[0, a)$  or  $[0, a]$ . When we know that  $g$  is strictly positive and bounded above by one on  $[0, a)$  or  $[0, a]$ , then zero is not an option and thus  $g(x) = e^{-\lambda x}$  for some  $0 \leq \lambda < \infty$ . Thus, for every  $z \in I_\theta$  and  $x \geq 0$  such that  $w = z + x \in \theta$  we have that for some  $0 \leq \lambda_\theta < \infty$

$$\frac{s_\theta(z + x)}{s_\theta(z)} = e^{-\lambda_\theta x} = \frac{e^{-\lambda_\theta(z+x)}}{e^{-\lambda_\theta z}} \quad (21)$$

which implies that for every  $z, w \in I_\theta$

$$s_\theta(w)e^{\lambda_\theta w} = s_\theta(z)e^{\lambda_\theta z} \equiv c_\theta, \quad (22)$$

as required. □

### 3 Maximum at a random time of a continuous time Markov process with no positive jumps

Consider a continuous time right continuous Markov process  $\{X(t)\}_{t \geq 0}$  with convex state space  $\mathfrak{X} \subset \mathbb{R}$ , having no positive jumps and with generator  $\mathcal{A}$ . As is customary, we denote  $\mathbb{P}_x$  and  $\mathbb{E}_x$  the distribution measure and the expected value when the process is initiated at  $x \in \mathfrak{X}$ . Assuming its existence, let  $f$  be strictly positive and nondecreasing function in the extended domain, which is bounded on  $(-\infty, x] \cap \mathfrak{X}$  for any  $x \in \mathbb{R}$  and for which

$$M(t) = f(X(t)) \exp\left(-\int_0^t \frac{\mathcal{A}f(X(s))}{f(X(s))} ds\right) \quad (23)$$

is a martingale with respect to the right continuous augmented filtration  $\{\mathcal{F}_t \mid t \geq 0\}$  generated by  $X$ . A sufficient condition for the latter is that  $f$  is bounded away from zero on  $\mathfrak{X}$  (e.g. [4], p.175). Furthermore we assume that  $\mathcal{A}f(x)$  is nonnegative for all  $x \in \mathfrak{X}$ . Now, denote  $\tau(y) = \inf\{t \mid X(t) > y\}$  (infinite if  $X$  never exceeds  $y$ ). Then  $\tau(y)$  is right



continuous in  $y$  and it is easy to check that  $\sup_{0 \leq s \leq t} X(s) \leq y$  if and only if  $\tau(y) \geq t$ . With  $a \wedge b = \min(a, b)$  it is well known that  $M(\tau(y) \wedge t)$  is also a martingale and moreover, our assumptions assure that it is also bounded. Finally, denoting  $\lambda(x) = \mathcal{A}f(x)/f(x)$ , then by the bounded convergence theorem we have that for  $y \geq x$  such that  $y \in \mathfrak{X}$ , if either  $\tau(y) < \infty$   $\mathbb{P}_x$ -a.s. (almost surely) or  $\int_0^\infty \lambda(X(s))ds = \infty$  on  $\{\tau(y) = \infty\}$  then

$$f(x) = \mathbb{E}_x M(0) = \mathbb{E}_x M(\tau(y)) = f(y) \mathbb{E}_x e^{-\int_0^{\tau(y)} \lambda(X(s))ds} . \quad (24)$$

In particular this means that it is impossible to find a positive  $f$  in the extended domain of  $\mathcal{A}$  such that  $\mathcal{A}f \geq 0$  and  $\int_0^{\tau(y)} \lambda(X(s))ds = \infty$   $\mathbb{P}_x$ -a.s. for some  $x$ .

Now, if we denote  $s(x) = 1/f(x)$ , we have that

$$\mathbb{E}_x e^{-\int_0^{\tau(y)} \lambda(X(s))ds} = \frac{s(y)}{s(x)} \quad (25)$$

for every  $y$  for which  $\tau(y) < \infty$   $\mathbb{P}_x$ -a.s. If  $Z$  is a random variable (possibly infinite) such that  $\mathbb{P}_x(Z > t | \mathcal{F}_t) = e^{-\int_0^t \lambda(X(s))ds}$ , then in fact

$$\mathbb{E}_x e^{-\int_0^{\tau(y)} \lambda(X(s))ds} = \mathbb{P}_x(Z > \tau(y)) = \mathbb{P}_x\left(\max_{0 \leq t \leq Z} X(t) > y\right). \quad (26)$$

Thus, we have that

$$P_x\left[\max_{0 \leq t \leq Z} X(t) > y\right] = \frac{s(y)}{s(x)} \quad (27)$$

for  $x \leq y$ . If one assumes that  $U$  is an independent Uniform(0,1) random variable (if there isn't then it is easy to artificially modify our probability space so that there is), then taking  $\mathcal{F}_t^U = \mathcal{F}_t \vee \sigma(U)$ , we see that  $M$  is a martingale also with respect to this new filtration. Thus, if we let  $F(X, t) = 1 - e^{-\int_0^t \lambda(X(s))ds}$  and  $G(X, u) = \inf\{t | F(X, t) \geq u\}$  then  $Z = G(X, U)$  has the correct conditional distribution.

### 3.1 Lévy processes

Sometimes, for various values of  $\alpha$ , we may be lucky to find a function  $f$  satisfying the above conditions and for which  $\lambda(x) = \alpha$ . In this case we immediately obtain the Laplace transform

$$\mathbb{E}_x e^{-\alpha \tau(y)} = \frac{s(y)}{s(x)}. \quad (28)$$

For a Lévy process with no positive jumps (in particular a Brownian motion) and

$$\varphi(\alpha) = \log \mathbb{E}_0 e^{\alpha X(1)} = c\alpha + \frac{\sigma^2}{2}\alpha^2 + \int_{(-\infty, 0)} (e^{\alpha y} - 1 - \alpha y 1_{(-1, 0)}(y)) \nu(dy) \quad (29)$$

then the equation that needs to be solved is the following

$$cf'(x) + \frac{\sigma^2}{2}f''(x) + \int_{(-\infty, 0)} (f(x+y) - f(x) - yf'(x) 1_{(-1, 0)}(y)) \nu(dy) = \alpha f(x) . \quad (30)$$

Fortunately, when  $X$  is not nonincreasing (the negative of a subordinator or the zero function), then  $\varphi$  has an inverse on  $[\beta, \infty)$  when  $\beta = \inf\{\alpha \mid \varphi(\alpha) > 0, \alpha > 0\}$ . It is well known that  $\beta = 0$  if  $\varphi'(0) \geq 0$  and  $\beta > 0$  otherwise. In this case, for every  $x \leq y$ ,  $\tau(y)$  is  $\mathbb{P}_x$ -a.s. finite and  $f(x) = e^{\varphi^{-1}(\alpha)x}$  for  $\alpha \geq \beta$  satisfies all the needed requirements and in particular solves (30). So, as is well known,

$$\mathbb{E}_x e^{-\alpha\tau(y)} = \mathbb{E}_x e^{-\alpha\tau(y)} \mathbb{1}_{\{\tau(y) < \infty\}} = \frac{s(y)}{s(x)} = \frac{f(x)}{f(y)} = e^{-\varphi^{-1}(\alpha)(y-x)}, \quad (31)$$

even if  $\tau(y)$  is not  $\mathbb{P}_x$ -a.s. finite (that is, when  $\varphi'(0) < 0$ ). In this particular case  $Z \sim \exp(\alpha)$  (independent of  $X$ ) and it follows as is also well known that

$$\max_{0 \leq t \leq Z} X(t) - X(0) \sim \exp(\varphi^{-1}(\alpha)),$$

so that this random variable obeys the standard (not-generalized) memoryless property.

### 3.2 Reflected Brownian motion

For the reflected Brownian motion on  $\mathfrak{X} = [0, \infty)$  with general drift, the generator is the same as the one for Brownian motion, only that its domain is reduced to twice differentiable functions for which  $f'(0) = 0$ . In this case one needs to compute  $\mu f'(x) + \frac{\sigma^2}{2} f''(x) = \alpha f(x)$  subject to  $f'(0) = 0$  for  $\alpha > 0$ . It is easy to check that any positive constant multiple of the function

$$f(x) = \frac{e^{a^+x}}{a^+} + \frac{e^{-a^-x}}{a^-} \quad (32)$$

where  $a^\pm = \frac{\sqrt{\mu^2 + 2\sigma^2\alpha} \pm \mu}{\sigma^2}$ , would do the trick. In particular it is positive and increasing due to  $0 < a^- < a^+$ . In this case it is well known that  $\tau(y) < \infty$   $\mathbb{P}_x$ -a.s. for any  $y > x$  regardless of the value of  $\mu$ , but it can also be inferred from this without resorting to anything else, by letting  $\alpha \rightarrow 0$ . Of course, this particular result is quite standard (e.g., problem 4 on p. 95 of [5]).

More generally of course, given a positive function  $\lambda$  if it is possible to find some  $f$  satisfying our assumptions for which  $\mathcal{A}f(x) = \lambda(x)f(x)$  for each  $x \in \mathfrak{X}$ , then if either  $y > x$  is such that  $\tau(y)$  is  $\mathbb{P}_x$ -a.s. finite or  $\lambda$  is bounded away from zero, then (27) is satisfied.

### 3.3 A growth collapse process with generalized memoryless jumps

In this section we consider a piecewise deterministic Markov process  $X_t$  with jumps that are governed by a jump measure with the generalized lack of memory property described above. See [2] and [8] for similar models.

Let  $\{X(t)\}_{t \geq 0}$  be a Markov process on  $\mathfrak{X} = [0, \infty)$  which is deterministically increasing with rate  $r(x)$  between randomly occurring downward jumps. More specifically, we assume that inbetween jumps  $dX_t = r(X_t) dt$ , that  $r(x)$  is positive and Lipschitz-continuous and that the time  $t^*(x, y) = \int_x^y 1/r(u) du$  that is needed to reach the level  $y$  from  $x$  in the absence of any jumps is finite for all  $x < y \in [0, \infty)$ . Let  $\kappa : \mathfrak{X} \rightarrow [0, \infty)$  denote the

state-dependent jump rate, i.e. if the process is in the state  $x \in \mathfrak{X}$ , then a jump occurs during the next  $\Delta t$  time units with probability  $\kappa(x)\Delta t + o(\Delta t)$  (and the probability to see more than one jump is  $o(\Delta t)$ ). We assume  $\kappa$  to be bounded. Given that there is a jump at time  $t$ , the process jumps from state  $x \in \mathfrak{X}$  into some measurable  $A \subset [0, x)$  with probability  $\nu(x, A)$ . We assume that for  $0 \leq y \leq x \leq z$  the kernel  $\nu$  has the special property that

$$\nu(z, y) = \nu(z, x)\nu(x, y) \quad (33)$$

holds (compare with (5)). Here we write  $\nu(x, y)$  for  $\nu(x, [0, y])$ . It is then easy to see that a similar situation as in Section 1 is present (let  $P(z, A) = \nu(-z, -A)$  for  $z \leq 0$  and  $A \subseteq (x, 0]$ ). It follows that there exists a family  $\{I_\theta \mid \theta \in \Theta\}$  of disjoint intervals and singletons with  $\cup_{\theta \in \Theta} I_\theta = [0, \infty)$  such that for each  $\theta$  for which  $I_\theta$  is not a singleton there exists a function  $s_\theta : I_\theta \rightarrow \mathbb{R}$  which is nonvanishing such that for every  $x, y \in I_\theta$  with  $x \leq y$  we have that

$$\nu(x, y) = \frac{s_\theta(y)}{s_\theta(x)}.$$

Note that  $s_\theta(y) : I_\theta \rightarrow [0, \infty)$  is nondecreasing and is not necessarily bounded. The infinitesimal generator of the Markov process  $X_t$  is given by

$$\mathcal{A}f(x) = r(x)f'(x) + \kappa(x) \int_0^x (f(y) - f(x))\nu(x, dy). \quad (34)$$

We assume that the domain  $\mathcal{D}_\mathcal{A}$  of  $\mathcal{A}$  consists of functions  $f$  that are absolutely continuous and for which the expectation of  $\sum_{0 < T_i \leq t} |f(X_{T_i-}) - f(X_{T_i})|$  is finite for every  $t \geq 0$ , where  $T_i$  denotes the  $i$ th jump time (see [3]).

The following Lemma generalizes formula (28) in [8].

**Lemma 2.** *Suppose that  $r(x)$ ,  $\kappa(x)$ ,  $\lambda(x)$  and  $s_\theta(x)$  are differentiable for  $x \in I_\theta$ . Define the functions  $a(x) = r'(x) + r(x)\xi(x) - \lambda(x) - \kappa(x)$  and  $b(x) = \lambda'(x) + \lambda(x)\xi(x)$ , where  $\xi(x) = \frac{s'_\theta(x)}{s_\theta(x)} - \frac{\kappa'(x)}{\kappa(x)}$  if  $\kappa(x) \neq 0$  and  $\xi(x) = 0$  otherwise. Any twice differentiable solution  $f$  with  $f'(x)s_\theta(x)$  being continuous of*

$$r(x)f''(x) + a(x)f'(x) - b(x)f(x) = 0, \quad (35)$$

fulfils  $\mathcal{A}f(x) = \lambda(x)f(x)$ .

*Proof.* The process  $X_t$ , if started in the state  $x \in I_\theta$  will leave  $I_\theta$  only at the moment when it passes through the upper boundary  $z^*(\theta)$  and  $\nu(x, y) = 0$  for  $y < z^*(\theta)$ . If  $x \in I_\theta$  we may hence write

$$\mathcal{A}f(x) = r(x)f'(x) + \frac{\kappa(x)}{s_\theta(x)} \int_{z^*(\theta)}^x \int_x^y f'(u) du s_\theta(dy), \quad x \in I_\theta.$$

Applying Fubini's theorem we can write this as

$$\mathcal{A}f(x) = r(x)f'(x) - \frac{\kappa(x)}{s_\theta(x)} \int_{z^*(\theta)}^x f'(u)s_\theta(u) du, \quad x \in I_\theta. \quad (36)$$

Then  $\mathcal{A}f(x) = \lambda(x)f(x)$  is equivalent to

$$\kappa(x) \int_{z_*(\theta)}^x f'(u)s_\theta(u) du = s_\theta(x) \left( r(x)f'(x) - \lambda(x)f(x) \right). \quad (37)$$

Differentiation yields

$$\begin{aligned} \frac{\kappa'(x)}{s_\theta(x)} \int_{z_*(\theta)}^x f'(u)s_\theta(u) du &= r(x)f''(x) + (r'(x) + r(x) \frac{s'_\theta(x)}{s_\theta(x)} \\ &\quad - \lambda(x) - \kappa(x))f'(x) - (\lambda'(x) + \lambda(x) \frac{s'_\theta(x)}{s_\theta(x)})f(x). \end{aligned}$$

If  $\kappa(x) \neq 0$  then we divide (37) by  $\kappa(x)$  and obtain (35) with  $\xi(x) = \frac{s'_\theta(x)}{s_\theta(x)} - \frac{\kappa'(x)}{\kappa(x)}$ . If  $\kappa(x) = 0$  then it follows from (37) that

$$r(x)f''(x) + (r'(x) - \lambda(x))f'(x) - \lambda'(x)f(x) = 0,$$

which is (35) with  $\xi(x) = 0$ .  $\square$

As is described earlier in the section via (27), the probability that the maximum process  $\max_{0 \leq t \leq Z} X(t)$  exceeds  $y$ , given  $X(0) = x$ , satisfies the generalized lack of memory property when  $Z$  is defined right before (26). More precisely,

**Corollary 1.** *Fix a  $\theta \in \Theta$  and suppose that  $f \in \mathcal{D}_\mathcal{A}$  is bounded away from zero (or is such that  $M(t)$  in (23) is a martingale) and solves equation (35) in  $I_\theta$ . Then*

$$P_x \left[ \max_{0 \leq t \leq Z} X(t) > y \right] = \frac{f(x)}{f(y)},$$

for all  $x, y \in I_\theta$  with  $x \leq y$ , where  $Z$  be a random variable, such that  $\mathbb{P}_x(Z > t | \mathcal{F}_t) = e^{-\int_0^t \lambda(X(s)) ds}$ .

In general (35) is not easy to solve and closed form solutions may be obtained only in certain cases. We provide two examples, where the coefficients  $a(x)$  and  $b(x)$  are such that a solution can be given.

**Example 1.** Equation (35) reduces to a differential equation with constant coefficients if

$$\begin{aligned} \frac{r'(x)}{r(x)} + \frac{s'_\theta(x)}{s_\theta(x)} - \frac{\kappa'(x)}{\kappa(x)} - \frac{\lambda(x) + \kappa(x)}{r(x)} &\equiv C \\ \text{and} \quad \frac{\lambda(x)}{r(x)} \left( \frac{\lambda'(x)}{\lambda(x)} + \frac{s'_\theta(x)}{s_\theta(x)} - \frac{\kappa'(x)}{\kappa(x)} \right) &\equiv D. \end{aligned}$$

For example suppose that  $\lambda(x) = c_1 e^{\alpha x}$ ,  $\kappa(x) = c_2 e^{\alpha x}$ ,  $r(x) = c_3 e^{\alpha x}$ ,  $s_\theta(x) = c_4 e^{\beta x}$ , with  $c_1, c_2, c_3, c_4, \beta \geq 0$  and  $\alpha \in \mathbb{R}$ . Then (35) reads

$$f''(x) + \left( \beta - \frac{c_1 + c_2}{c_3} \right) f'(x) - \frac{c_1 \beta}{c_3} f(x) = 0,$$

which is solved by  $f(x) = Ae^{a^-x} + Be^{a^+x}$ , where

$$a^\pm = \frac{1}{2} \left( \beta - \frac{c_1 + c_2}{c_3} \pm \sqrt{\left( \beta - \frac{c_1 + c_2}{c_3} \right)^2 + 4 \frac{c_1 \beta}{c_3}} \right).$$

If we set  $f(z_*(\theta)) = 1$  (w.l.o.g.), then  $f'(z_*(\theta)) = \lambda(z_*(\theta))/r(z_*(\theta)) = c_1/c_3$ . This leads to the final solution

$$f(x) = \frac{a^+ - \frac{c_1}{c_3}}{a^+ - a^-} e^{a^-(x-z_*(\theta))} + \frac{\frac{c_1}{c_3} - a^-}{a^+ - a^-} e^{a^+(x-z_*(\theta))}.$$

**Example 2.** This example is a generalization of Example (A), Section 4.1 in [8]. Suppose that the jump measure  $\nu(x, y) = s_\theta(y)/s_\theta(x)$  is defined such that for some  $\alpha > 0$   $s_\theta(x)\lambda(x) = \alpha\kappa(x)$ . Then  $\xi(x) = -\lambda'(x)/\lambda(x)$  and as a consequence the second coefficient  $b(x)$  is zero (while  $a(x) = r'(x) - r(x)\lambda'(x)/\lambda(x) - \lambda(x) - \kappa(x)$ ). Hence (35) becomes

$$r(x)f''(x) + a(x)f'(x) = 0, \quad (38)$$

which is solved by

$$f(x) = f(z_*(\theta)) + f'(z_*(\theta)) \frac{r(z_*(\theta))}{\lambda(z_*(\theta))} \int_{z_*(\theta)}^x \frac{\lambda(u)}{r(u)} e^{\int_{z_*(\theta)}^u \frac{\lambda(w)+\kappa(w)}{r(w)} dw} du.$$

Note that since  $\mathcal{A}f(x) = \lambda(x)f(x)$  it follows that  $\lambda(z_*(\theta))f(z_*(\theta)) = r(z_*(\theta))f'(z_*(\theta))$  and hence, choosing w.l.o.g.  $f(z_*(\theta)) = 1$ , we obtain the solution

$$f(x) = 1 + \int_{z_*(\theta)}^x \frac{\lambda(u)}{r(u)} e^{\int_{z_*(\theta)}^u \frac{\lambda(w)+\kappa(w)}{r(w)} dw} du.$$

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