# On a pairing heuristic in binpacking 

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## Memorandum COSOR 86-13

On a pairing heuristic in binpacking by

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Eindhoven, the Netherlands
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## ON A PAIRING HEURISTIC IN BINPACKING


#### Abstract

For the analysis of a pairing heuristic in binpacking an important result is used without proof in [1] and [2].

In this note we discuss this result and give a detailed proof of it.


## Introduction

Let $n \in \mathbb{N}$ be given and suppose ( $\underline{X}_{1}, \cdots, \underline{X}_{n}$ ) is a $n$-dimensional stochastic vector with joint density $f\left(x_{1}, \cdots, x_{n}\right)$

Moreover assume
(i) $0 \leq \underline{X}_{i} \leq 1 \quad i=1, \cdots, n$
(ii) The stochastic vector $\left(\underline{X}_{\sigma(1)}, \underline{X}_{\sigma(2)}, \cdots, \underline{X}_{\sigma(n)}\right)$ is distributed as $\left(\underline{X}_{1}, \underline{X}_{2}, \cdots, \underline{X}_{n}\right)$ for every permutation $\sigma$ on $\{1, \cdots, n\}$.
(iii) $f\left(x_{1}, x_{2}, \cdots, x_{n}\right)=f\left(1-x_{1}, \cdots, x_{n}\right)$

## Remark

Condition (ii) states that we are dealing with a finite sequence of so-called exchangeable random variables (cf. [3]), while condition (iii) is a symmetry condition.

Note that by (ii) the symmetry in (iii) holds in every component.

Before stating the main result introduce the following notations

$$
\begin{aligned}
& 1_{A}:=\left\{\begin{array}{lc}
1 & \text { if the event } A \text { happens } \\
0 & \text { otherwise }
\end{array}\right. \\
& \underline{Y}_{i}:=\left(1-\underline{X}_{i}\right) 1_{\left(X_{i}>\frac{1}{2}\right)}+\underline{X}_{i} \quad 1_{\left(X_{i} \leq \frac{1}{2}\right)} . \quad i=1, \cdots, n \\
& (\underline{i}):=\left\{\begin{array}{ll}
+1 & \text { if } \underline{X}_{i}>\frac{1}{2} \\
-1 & \text { if } \underline{X}_{i} \leq \frac{1}{2}
\end{array} \quad i=1, \cdots, n\right.
\end{aligned}
$$

If we order the random variables $\underline{Y}_{i}$ in non-decreasing order, say $\underline{Y}_{i_{1}} \leq \underline{Y}_{i_{2}} \leq \cdots \leq \underline{Y}_{i_{n}}$, we denote by $\left(\underline{i_{k}}\right)$ the label of the $k$-order statistic of the sequence $\left\{\underline{Y}_{i}\right\}_{i=1}^{n}$.
Now the main result reads as follows.

## Theorem 1

Suppose the random variables $\left\{\underline{X}_{i}\right\}_{i=1}^{n}$ satisfy the conditions (i), (ii) and (iii).
Then the following results hold
a) $\quad\left(\underline{Y}_{i}, i \in A\right\}$ and $\{(i), i \in A\}$ are independent for every subset $A \subset\{1,2, \cdots, n\}$
b) $\mathbb{P}\left\{\left(\underline{i_{k}}\right)=\pi\left(i_{k}\right), k \in A\right\}=\prod_{k \in A} \mathbb{P}\left\{\left(\underline{i_{k}}\right)=\pi\left(i_{k}\right)\right\}$

$$
=2^{-|A|}
$$

for every subset $A \subset\{1,2, \cdots, n\}$ and
for every function $\pi:\{1,2, \cdots, n\} \rightarrow\{-1,1\}$.

Proof For every sequence $\left\{y_{i}\right\}_{i=1}^{n}$ with $y_{i} \in\left(0, \frac{1}{2}\right)$ and $\sigma$ some permutation on $\{1, \cdots, n\}$ we obtain

$$
\begin{aligned}
& \left.\mathbb{P}\left\{\underline{Y}_{\sigma(i)} \leq y_{\sigma(i)}, \underline{(\sigma(i)}\right)=\pi(\sigma(i)), i=1, \cdots, k\right\}= \\
& =\mathbb{P}\left\{1-\underline{X}_{\sigma(i)} \leq y_{\sigma(i)}(i \in C) \wedge \underline{X}_{\sigma(i)} \leq y_{\sigma(i)}(i \in\{1, \cdots, k\}-C)\right\}
\end{aligned}
$$

where $1 \leq k \leq n$ and $C:=\{j: 1 \leq j \leq k \& \pi(\sigma(j))=1\}$
By (ii) and (iii) it follows easily

$$
\mathbb{P}\left\{\underline{Y}_{\sigma(i)} \leq y_{\sigma(i)},(\underline{\sigma(i)})=\pi(\sigma(i)) i=1, \cdots, k\right\}=
$$

(1) $\mathbb{P}\left\{\underline{X}_{\sigma(i)} \leq y_{\sigma(i)} ; i=1, \cdots, k\right\}=\mathbb{P}\left\{\underline{X}_{i} \leq y_{\sigma(i)} ; i=1, \cdots, k\right\}$
and this implies

$$
\begin{aligned}
& \quad \mathbb{P}\left\{\underline{Y}_{\sigma(i)} \leq y_{\sigma(i)} ; i=1, \cdots, k\right\}= \\
& =\sum_{\tau \in D} \mathbb{P}\left\{\underline{Y}_{\sigma(i)} \leq y_{\sigma(i)},(\underline{\sigma(i)})=\tau(\sigma(i)) ; i=1, \cdots, k\right\}= \\
& \text { (2) }=\sum_{\tau \in D} \mathbb{P}\left\{\underline{X}_{i} \leq y_{\sigma(i)} ; i=1, \cdots, k\right\}=2^{k} \mathbb{P}\left\{\underline{X}_{i} \leq y_{\sigma(i)} ; i=1, \cdots, k\right\}
\end{aligned}
$$

where $D$ is the set of functions $\tau:\{1,2, \cdots, n\} \rightarrow\{-1,+1\}$ which are different on $\{\sigma(1), \cdots, \sigma(k)\}$. Moreover by (1)

$$
\begin{aligned}
& \quad \mathbb{P}\{(\underline{\sigma(i))}=\pi(\sigma(i)) ; i=1, \cdots, k\}= \\
& \left.=\mathbb{P}\left(\underline{Y}_{\sigma(i)} \leq \frac{1}{2}, \underline{(\sigma(i)}\right)=\pi(\sigma(i)), i=1, \cdots, k\right\}= \\
& \text { (3) } \quad=\mathbb{P}\left\{\underline{X}_{i} \leq \frac{1}{2} ; i=1, \cdots, k\right\} .
\end{aligned}
$$

Since the density $f\left(x_{1}, \cdots, x_{n}\right)$ is symmetric it is easy to prove that for every $1 \leq l \leq n-1$

$$
\mathbb{P}\left\{\underline{X}_{1} \leq \frac{1}{2}, \cdots, \underline{X}_{l} \leq \frac{1}{2}\right\}=2 \mathbb{P}\left\{\underline{X}_{1} \leq \frac{1}{2}, \cdots, \underline{X}_{l+1} \leq \frac{1}{2}\right\}
$$

and this implies using $\mathbb{P}\left\{\underline{X}_{1} \leq \frac{1}{2}\right\}=\frac{1}{2}$ that
(4) $\mathbb{P}\left\{\underline{X}_{1} \leq \frac{1}{2}, \cdots \underline{X}_{l} \leq \frac{1}{2}\right\}=2^{-l}$

Now by the relations (1), (2), (3) and (4)

$$
\begin{aligned}
& \left.\mathbb{P}\left\{\underline{Y}_{\sigma(i)} \leq y_{\sigma(i)}, \underline{(\sigma(i)}\right)=\pi(\sigma(i)) ; i=1, \cdots, k\right\}= \\
& \mathbb{P}\left\{\underline{X}_{i} \leq y_{\sigma(i)} ; i=1, \cdots, k\right\}= \\
& 2^{-k} .2^{k} \mathbb{P}\left(\underline{X}_{i} \leq y_{\sigma(i)} ; i=1, \cdots, k\right\}= \\
& \mathbb{P}\{\underline{(\sigma(i))}=\pi(\sigma(i)) ; i=1, \cdots, k\} . \mathbb{P}\left\{\underline{Y}_{\sigma(i)} \leq y_{\sigma(i)} ; i=1, \cdots, k\right\}
\end{aligned}
$$

and so we have proved the result in (a)

In order to prove the result in b) we note that for every subset $A \subset\{1,2, \cdots, n\}$ and every function $\pi:(1,2, \cdots, n\} \rightarrow\{-1,+1\}$

$$
\begin{aligned}
& \mathbb{P}\left\{\left(\underline{i_{k}}\right)=\pi\left(i_{k}\right) ; k \in A\right\}= \\
& =\sum_{\sigma} \mathbb{P}\left\{\underline{Y}_{\sigma(i)} \leq \frac{1}{2}, \underline{Y}_{\sigma(1)} \leq \underline{Y}_{\sigma(2)} \leq \cdots \leq \underline{Y}_{\sigma(n)},(\underline{\sigma(k)})=\pi(\sigma(k)) ; k \in A\right\} \\
& =\sum_{\sigma} \mathbb{P}\left\{\underline{Y}_{\sigma(i)} \leq \frac{1}{2}, \underline{Y}_{\sigma(1)} \leq \cdots \leq \underline{Y}_{\sigma(n)}\right\} \mathbb{P}\{(\underline{\sigma(k)})=\pi(\sigma(k)) ; k \in A\}
\end{aligned}
$$

where we have used (a) to obtain the last equality.
Hence

$$
\mathbb{P}\left\{\left(\underline{i_{k}}\right)=\pi\left(i_{k}\right) ; k \in A\right\}=
$$

$$
\begin{aligned}
& =2^{-|A|} \sum_{\sigma} \mathbb{P}\left\{\underline{Y}_{\sigma(1)} \leq \cdots \leq \underline{Y}_{\sigma(n)}, \underline{Y}_{\sigma(i)} \leq \frac{1}{2} ; i=1, \cdots, n\right\} \\
& =2^{-|A|}=\prod_{k \in A} \mathbb{P}\left\{\left(\underline{i}_{\underline{k}}\right)=\pi\left(i_{k}\right)\right\}
\end{aligned}
$$

## References

[1] Csirik, J., Frenk, J.B.G., Galambos, G., Rinnooy Kan, A.H.G., A probabilistic analysis of the dual bin packing problem, to appear.
[2] Karp, R.M., Lecture Notes (unpublished), Computer Science Department, University of California, Berkeley, 1984.
[3] Feller, W., An Introduction to Probability Theory and its Applications, vol II, Wiley, 1971.

