Syndrome decoding of convolutional codes

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SYNDROME DECODING OF CONVOLUTIONAL CODES

by

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Abstract

The classical Viterbi decoder recursively finds the trellis path (codeword) closest to the received data. Given the received data the syndrome decoder first forms a syndrome, instead. Having found the syndrome, that only depends on the channel noise, a recursive algorithm like Viterbi's determines the noise sequence of minimum Hamming weight that can be a possible cause of this syndrome. Given the estimate of the noise sequence one derives an estimate of the original data sequence. Whereas, the bit error probability of the syndrome decoder is no different from that of the classical Viterbi decoder, the syndrome decoder can be naturally implemented using a read only memory (ROM), thus obtaining a considerable saving in hardware.

I. INTRODUCTION

The principle of syndrome decoding of convolutional codes will be explained using the binary code generated by the encoder of Fig. 1.

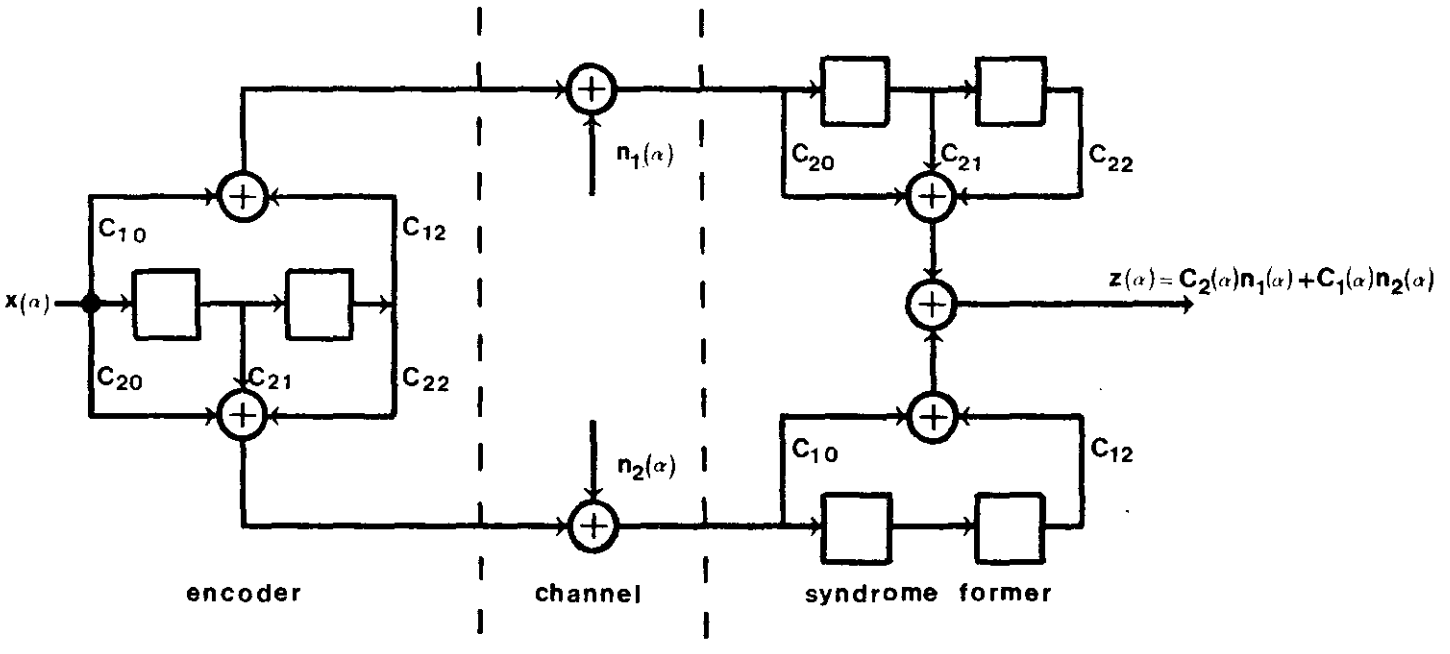

Fig. 1. Encoding and syndrome forming for a R=½ code

The additions in Fig. 1 are modulo-2, and all binary sequences $b_0, b_1, b_2, \ldots$ are represented as power series $b(\alpha)=b_0+b_1\alpha+b_2\alpha^2+\ldots$. The encoder has connection polynomials $C_1(\alpha)= 1+\alpha^2$, and $C_2(\alpha)= 1+\alpha+\alpha^2$. Hence, the encoder outputs are $C_1(\alpha)x(\alpha)$, and $C_2(\alpha)x(\alpha)$. The syndrome $z(\alpha)$ only depends on $n_1(\alpha)$ and $n_2(\alpha)$, i.e. not on the data sequence $x(\alpha)$, for

$$z(\alpha) = C_2(\alpha)[C_1(\alpha)x(\alpha)+n_1(\alpha)]+C_1(\alpha)[C_2(\alpha)x(\alpha)+n_2(\alpha)] =$$

$$= C_2(\alpha)n_1(\alpha)+C_1(\alpha)n_2(\alpha) \qquad (1)$$

Having formed the syndrome $z(\alpha)$, the next section describes a recursive algorithm like Viterbi's [1] to determine the noise sequence pair $[\hat{n}_1(\alpha), \hat{n}_2(\alpha)]$ of minimum Hamming weight that can be a possible cause of this syndrome.

Given the estimate $[\hat{n}_1(\alpha), \hat{n}_2(\alpha)]$ of the noise sequence pair one derives an estimate $\hat{x}(\alpha)$ of the original data sequence $x(\alpha)$ as follows. For a noncatastrophic code $C_1(\alpha)$ and $C_2(\alpha)$ are relatively prime. Hence, by Euclids algorithm [2] there exist polynomials $d_1(\alpha)$ and $d_2(\alpha)$ such that $d_1(\alpha)C_1(\alpha)+d_2(\alpha)C_2(\alpha)=1$. For the example of Fig. 1 we have $d_1(\alpha)=1+\alpha$, $d_2(\alpha)=\alpha$. We receive the sequence pair

$$y_i(\alpha) = C_i(\alpha)x(\alpha)+n_i(\alpha) \quad ; \quad i=1,2 \quad , \qquad (2)$$

and from the estimate

$$\hat{x}(\alpha) = d_1(\alpha)[y_1(\alpha)+\hat{n}_1(\alpha)]+d_2(\alpha)[y_2(\alpha)+\hat{n}_2(\alpha)] \qquad (3)$$

Note that if the noise sequence estimate $[\hat{n}_1(\alpha), \hat{n}_2(\alpha)]$ is correct we have

$$y_i(\alpha)+\hat{n}_i(\alpha) = C_i(\alpha)x(\alpha)+n_i(\alpha)+\hat{n}_i(\alpha) = C_i(\alpha)x(\alpha) \quad ; \quad i=1,2 \quad ,$$

and, hence,

$$\hat{x}(\alpha) = d_1(\alpha)C_1(\alpha)x(\alpha)+d_2(\alpha)C_2(\alpha)x(\alpha) = x(\alpha)$$

Note that (3) for the estimate $\hat{x}(\alpha)$ of the data sequence $x(\alpha)$ can be rewritten as

$$\hat{x}(\alpha) = [d_1(\alpha)y_1(\alpha)+d_2(\alpha)y_2(\alpha)]+\omega(\alpha), \qquad (4)$$

where

$$\omega(\alpha) = d_1(\alpha)\hat{n}_1(\alpha)+d_2(\alpha)\hat{n}_2(\alpha) \qquad (5)$$

The term in square brackets in (4) can be computed directly from the received data using very simple circuitry. As there is no need to distinguish between pairs $[\hat{n}_1(\alpha), \hat{n}_2(\alpha)]$, and $[\hat{n}_1(\alpha), \hat{n}_2(\alpha)]'$ that lead to the same value for $\omega(\alpha)$ in (5), the algorithm to be discussed in the next section computes $\omega(\alpha)$ directly.

II. THE ALGORITHM

In Fig. 2 we have redrawn the syndrome former. As, according to (1),

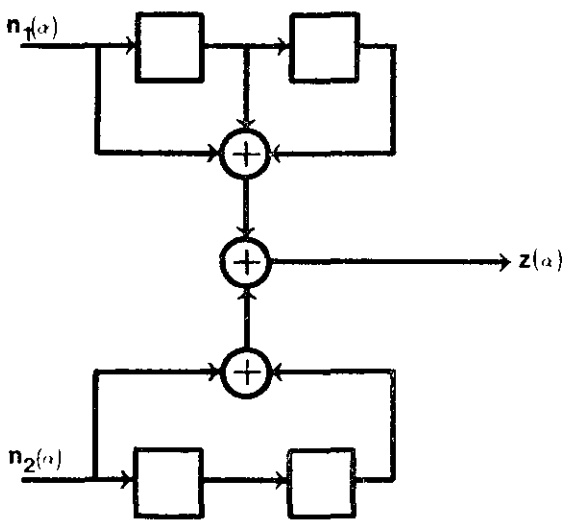

Fig. 2. The syndrome former

the syndrome $z(\alpha)$ only depends on the noise pair $[n_1(\alpha), n_2(\alpha)]$ all other binary sequences have been omitted from Fig. 2. For minimum distance decoding we are now presented with the following problem. Given the syndrome $z(\alpha)$ determine the noise pair $[\hat{n}_1(\alpha), \hat{n}_2(\alpha)]$ of minimum Hamming weight that can be a cause of this syndrome.

At first sight the state diagram of the syndrome former of Fig. 2 has $2^4 = 16$ states and, hence, is more complicated than the state diagram used to implement the classical Viterbi decoder [1] that has only $2^2 = 4$ states. However, a closer inspection of Fig. 2 reveals that the syndrome former has also $2^2 = 4$ states. In general, for an encoder with ν memory stages the syndrome former has 2^ν states just like the state diagram used to implement the classical Viterbi decoder. This can be seen as follows. Writing

$$[n_1(\alpha),\ n_2(\alpha)] = [n_{10},\ n_{20}] + [n_{11},\ n_{21}]\alpha + [n_{12},\ n_{22}]\alpha^2 + \ldots\ , \quad (6)$$

each successive binary coefficient pair $[n_{1k},\ n_{2k}]$, k=0,1,2,... , can be arbitrarily replaced by its modulo-2 complement $[\bar{n}_{1k},\ \bar{n}_{2k}]$ without altering the syndrome $z(\alpha)$. Hence, of the $2^{2\nu}$ different memory contents of Fig. 2, 2^{ν} are equivalent as far as $z(\alpha)$ is concerned leaving $2^{2\nu}/2^{\nu}=2^{\nu}$ different states. Fig. 3 gives the state diagram of the syndrome former of Fig. 2. Solid transitions in Fig. 3 correspond to $z_k=0$ and dashed transitions

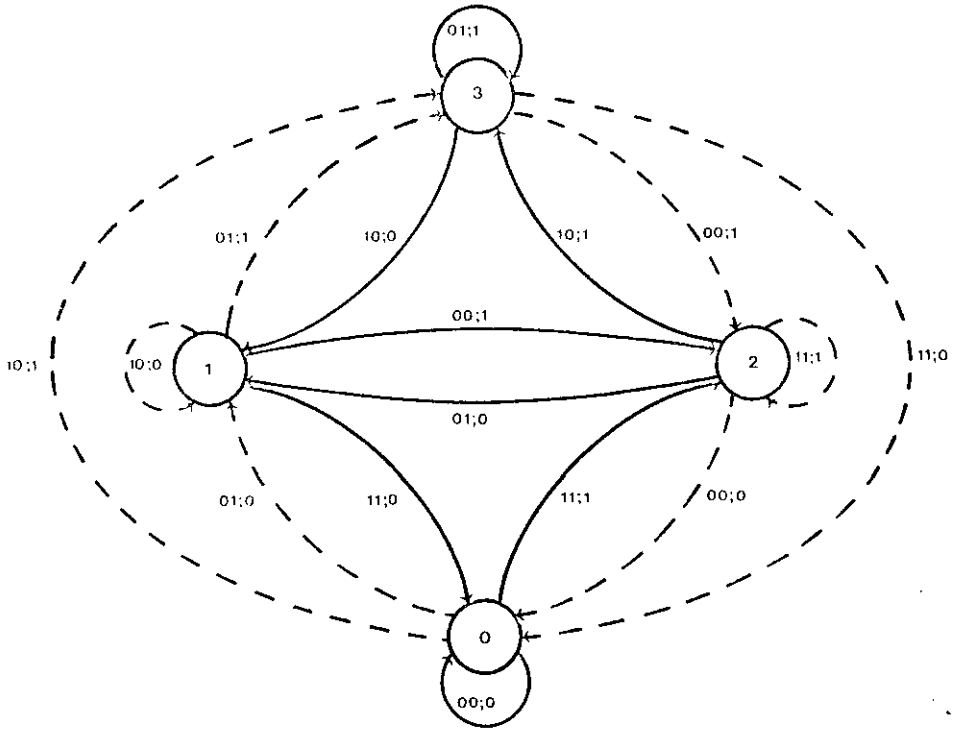

Fig. 3. State diagram of syndrome former

to $z_k=1$, k=0,1,2,... . Next to each transition one finds the value of $\hat{n}_{k1}$, $\hat{n}_{k2}$; ω_k , k=0,1,2,... . Fig. 4 gives the k-th, k=0,1,2,... , section of the trellis diagram that corresponds to the state diagram of Fig.3. The algorithm that determines $\omega(\alpha)$ according to (5) now operates as follows. With each state in Fig. 4 we associate a metric $M_j(k)$, j=0,1,2,3 , k=0,1,2,... , that equals the minimum Hamming weight of a path, $[\hat{n}_1(\alpha),\ \hat{n}_2(\alpha)]^{(j)}$, leading from state j=0 at time k=0 to that particular

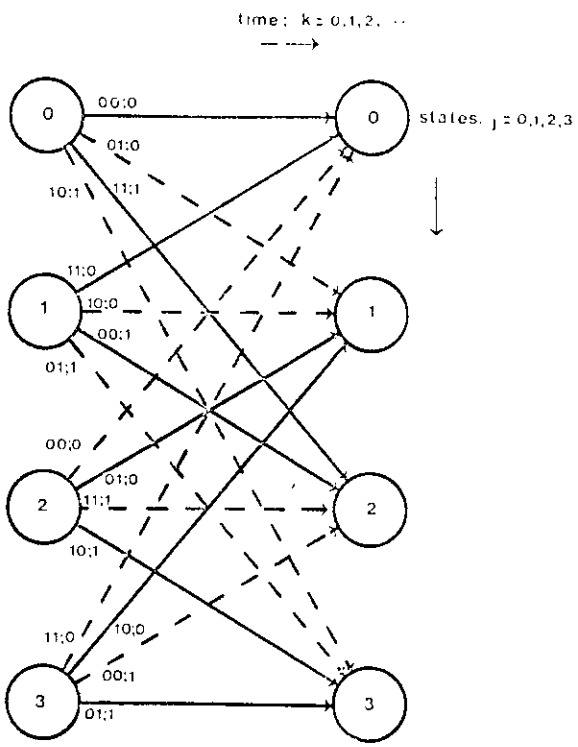

Fig. 4. The k-th section of the trellis diagram, k=0,1,2,...

state. This path has a solid or a dashed ℓ-th branch, $0 \le \ell \le k-1$, according to whether $z_\ell = 0$ or $z_\ell = 1$, respectively. The metric $M_j(k+1)$ at time k+1 can be determined recursively, i.e.

$$M_0(k+1) = \bar{z}_k \min [M_0(k), M_1(k)+2] + z_k \min [M_2(k), M_3(k)+2] \quad (7a)$$

$$M_1(k+1) = \bar{z}_k \min [M_2(k)+1, M_3(k)+1] + z_k \min [M_0(k)+1, M_1(k)+1] \quad (7b)$$

$$M_2(k+1) = \bar{z}_k \min [M_0(k)+2, M_1(k)] + z_k \min [M_2(k)+2, M_3(k)] \quad (7c)$$

$$M_3(k+1) = \bar{z}_k \min [M_2(k)+1, M_3(k)+1] + z_k \min [M_0(k)+1, M_1(k)+1] \quad (7d)$$

Given the value of z_k, i.e. $z_k = 0$ or $z_k = 1$, each (k+1)-state can be reached from two k-states. For each of these two k-states add to the metric, the Hamming weight of the transition, i.e. of $[\hat{n}_{k1}, \hat{n}_{k2}]$, to the particular (k+1)-state. The minimum of the two values thus obtained is $M_j(k+1)$. The

transition associated with the minimum value is called the "survivor". In case of a tie, choose the survivor at random among the two candidates. The survivor for (k+1)-state j=0,1,2,3 can be specified by the associated k-state $j_j(k)$=0,1,2,3. Going back from a (k+1)-state each time choosing the survivor we obtain the path, $[\hat{n}_1(\alpha), \hat{n}_2(\alpha)]^{(j)}$, j=0,1,2,3 , of minimum Hamming weight leading to that particular (k+1)-state. The coefficients $\omega^{(j)}_{k-D+1}$, $\omega^{(j)}_{k-D+2}$, ..., $\omega^{(j)}_{k}$, associated with the path, $[\hat{n}_1(\alpha), \hat{n}_2(\alpha)]^{(j)}$, of minimum Hamming weight are stored in the path register for the j-th state, j=0,1,2,3. If

$$M_{j_0}(k+1) = \min_{j} M_j(k+1) \tag{8}$$

we set

$$\omega_{k-D+1} = \omega^{(j_0)}_{k-D+1} \tag{9}$$

If more than one j_0 satisfies (8) we make an arbitrary selection among the candidates. The longer the path register length D the smaller the resulting bit error probability, P_b. Increasing D beyond 5(ν+1) does not lead to an appreciable further decrease in P_b. We have done detailed calculations concerning the relationship between D and P_b, which will be published shortly. The next section is concerned with the practical implementation of the syndrome decoder.

III. IMPLEMENTATION

Using (7) we construct Table I. The first column just numbers the rows of the table. The second column lists all possible metric combinations $M_0(k)$, $M_1(k)$, $M_2(k)$, $M_3(k)$ at time k. As only the differences between the metrics of a quadruple matter we subtract from each member of a quadruple of metrics the minimum value of the quadruple, i.e. all quadruples of metrics in Table I have one or more zeros. Column 3 and 4 apply to the case that $z_k=0$ and columns 5 and 6 to the case that $z_k=1$. Columns 3 and 5 list the survivors $j_0(k)$, $j_1(k)$, $j_2(k)$, $j_3(k)$, and columns 4 and 6 the new metrics $M_0(k+1)$, $M_1(k+1)$, $M_2(k+1)$, $M_3(k+1)$ as given by (7). If there is a choice of survivors the candidates are placed within parentheses in the survivor columns.

row number	old metrics	$z_k=0$		$z_k=1$	
		survivors	new metrics	survivors	new metrics
0	0000	0(2,3) 1 (2,3)	0101	2 (0,1)3(0,1)	0101
1	0101	0 2 1 2	0111	2 0 3 0	0111
2	0111	0(2,3) 1 (2,3)	0212	2 0 3 0	0000
3	0212	0 2 (0,1) 2	0222	2 0 3 0	0010
4	0222	0(2,3)(0,1)(2,3)	0323	2 0 3 0	1010
5	0010	0 3 1 3	0101	2 (0,1)3(0,1)	1101
6	0323	0 2 0 2	0323	2 0 3 0	1020
7	1010	0 3 1 3	1101	2 1 3 1	1101
8	1101	0 2 1 2	0000	2 (0,1)3(0,1)	0212
9	1020	0 3 1 3	1101	(2,3) 1 3 1	2101
10	2101	0 2 1 2	1000	2 1 3 1	0212
11	1000	0(2,3) 1 (2,3)	1101	2 1 3 1	0101

TABLE I. Metric transitions

Table I contains more information than is necessary for the actual implementation of the syndrome decoder. As explained in section II knowledge of the successive survivors for each state, together with the index j_0 of the minimum within each new quadruple of metrics suffices to determine the key sequence $\omega(\alpha)$ of (5). Hence, we omit the quadruples of metrics from Table I and store the resulting Table II in a ROM. The

old ROM-address	$z_k=0$			$z_k=1$		
	survivors	new ROM-address	index j_0	survivors	new ROM-address	index j_0
0	0 (2,3) 1 (2,3)	1	(0,2)	2 (0,1) 3 (0,1)	1	(0,2)
1	0 2 1 2	2	0	2 0 3 0	2	0
2	0 (2,3) 1 (2,3)	3	0	2 0 3 0	0	(0,1,2,3)
3	0 2 (0,1) 2	4	0	2 0 3 0	5	(0,1,3)
4	0 (2,3) (0,1) (2,3)	6	0	2 0 3 0	7	(1,3)
5	0 3 1 3	1	(0,2)	2 (0,1) 3 (0,1)	8	2
6	0 2 0 2	6	0	2 0 3 0	9	(1,3)
7	0 3 1 3	8	2	2 1 3 1	8	2
8	0 2 1 2	0	(0,1,2,3)	2 (0,1) 3 (0,1)	3	0
9	0 3 1 3	8	2	(2,3) 1 3 1	10	2
10	0 2 1 2	11	(1,2,3)	2 1 3 1	3	0
11	0 (2,3) 1 (2,3)	8	2	2 1 3 1	1	(0,2)

TABLE II. Contents of the ROM

operation of the core part of the syndrome decoder can now be explained using the block diagram of Fig. 5. Assume that at time k the ROM address register, AR, contains (AR)=7 and the ROM data register, DR, contains (DR)=(ROM,7). Let $z_k=1$. Note, see Fig. 4, that $\omega_k^{(0)}=\omega_k^{(1)}=0$, $\omega_k^{(2)}=\omega_k^{(3)}=1$ independent of k=0,1,2,... , i.e. always fill the left

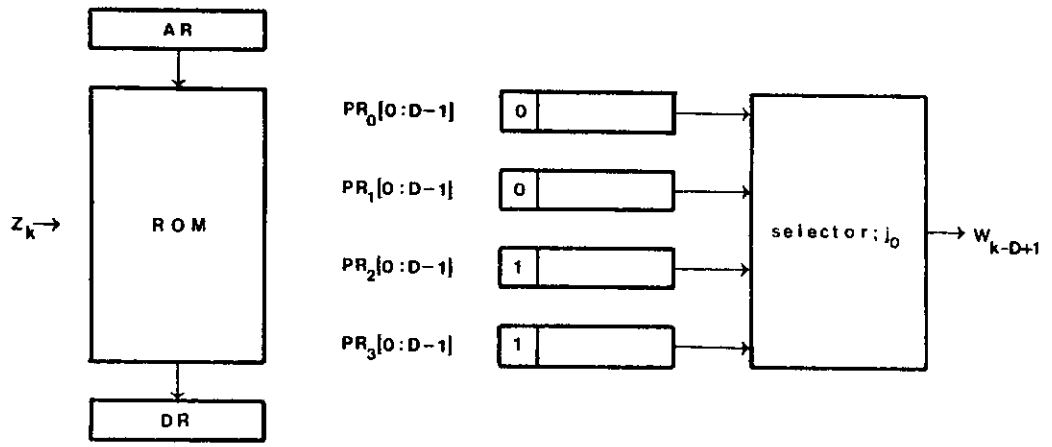

Fig. 5. Block diagram of the core of the syndrome decoder

most stages of the four path registers, $PR_0[0:0]$, $PR_1[0:0]$, $PR_2[0:0]$, $PR_3[0:0]$, with 0011, respectively. Then according to row 7 and column 5 of Table II, or according to the contents, (DR), of the DR, replace

$PR_0[1:D-1] \leftarrow$CONTENTS $PR_2[1:D-1]$

$PR_1[1:D-1] \leftarrow$CONTENTS $PR_1[1:D-1]$

$PR_2[1:D-1] \leftarrow$CONTENTS $PR_3[1:D-1]$

$PR_3[1:D-1] \leftarrow$CONTENTS $PR_1[1:D-1]$.

The right most digit, $PR_0[D-1:D-1]$, $PR_1[D-1:D-1]$, $PR_2[D-1:D-1]$, $PR_3[D-1:D-1]$, of all four path registers is fed to the selector, see Fig. 5, that determines ω_{k-D+1} according to (9) using the entry in row 7 and column 7, i.e. $j_0=2$, of Table II which can also be found in the DR. To complete the k-th cycle of the syndrome decoder, set (AR)=8 and read DR$\leftarrow$(ROM,8).

The ROM-decoder for the code of Fig. 1 has been realized in hardware using path registers of length D=11. The solid line in Fig. 6 gives the

measured bit error probability, P_b, as a function of the transition probability, p, of the binary symmetric channel. The dashed curve is an upper bound [3] on the bit error probability, P_b.

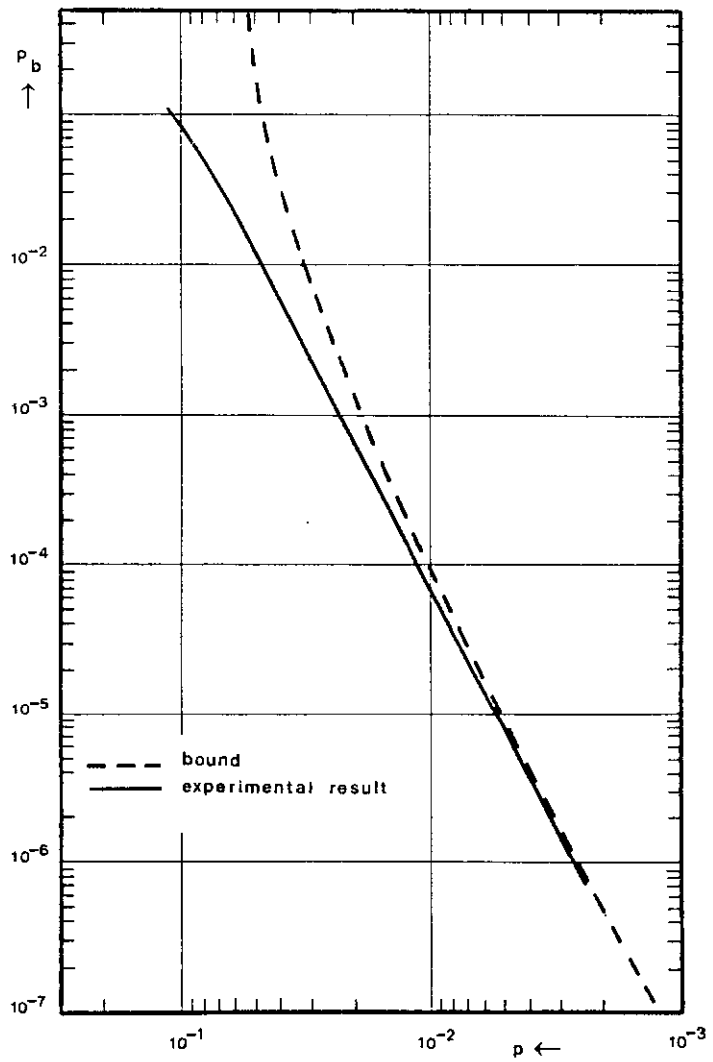

Fig. 6. Bit error rate P_b versus channel transition probability p.

IV. CONCLUSIONS

This paper describes a syndrome decoder for convolutional codes. The recursive algorithm that forms the core part of the decoder can be naturally implemented with a ROM. Using the same type of I.C.'s the syndrome decoder requires less than one third of the hardware that is necessary to implement the classical Viterbi decoder. A program has been developed that computes the contents of the ROM for an arbitrary rate $\frac{1}{2}$ binary convolutional code. This program enables us to quickly design an extremely efficient minimum distance decoder.

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