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Note on the approximation of distributions on \mathbb{Z}_+ by mixtures of negative binomial distributions

F.W. Steutel and M.J.A. van Eenige

Abstract

It is shown that the distributions on \mathbb{Z}_+ that can be approximated by mixtures of negative binomial distributions, are precisely the so-called Poisson mixtures, i.e., mixtures of Poisson distributions.

1 Introduction

It is well known that every distribution on \mathbb{R}_+ is the weak limit of a sequence of distributions with rational Laplace-Stieltjes transforms (LSt's), in fact, of mixtures (with positive weights) of Gamma distributions. More precisely, if \hat{F} is the LSt of a distribution function (df) on \mathbb{R}_+ , then there are $p_{k,n} > 0$, $\lambda_{k,n} \in (0, \infty)$ and $r_{k,n} \in \mathbb{N}$ such that

$$\hat{F}(s) = \lim_{n \to \infty} \sum_{k=1}^{n} p_{k,n} \left(\frac{\lambda_{k,n}}{\lambda_{k,n} + s} \right)^{r_{k,n}} \tag{1}$$

(see e.g. p. 32 in Schassberger (1973) and pp. 78-79 in Neuts (1981)). An easy proof of (1) consists of the following two facts.

- (i) Every df on ℝ₊ can be approximated by step functions, i.e., by mixtures of degenerate df's.
- (*ii*) A degenerate df concentrated at c can be approximated by a Gamma distribution: the law of large numbers implies that

$$\frac{c}{n}(X_1+\cdots+X_n)\overset{w}{\longrightarrow}c,$$

where X_1, \dots, X_n are i.i.d. and exponentially distributed with mean 1.

A natural question to ask is, 'What distributions on \mathbb{Z}_+ can be approximated by mixtures of negative binomial distributions, the analogues of Gamma distributions?'. This question will be answered in the next section.

2 Distributions on \mathbb{Z}_+

The analogue of (1) for the generating function *P* of a distribution on \mathbb{Z}_+ would be

$$P(z) = \lim_{n \to \infty} \sum_{k=1}^{n} p_{k,n} \left(\frac{1 - q_{k,n}}{1 - q_{k,n} z} \right)^{r_{k,n}},$$
(2)

with $p_{k,n} > 0$, $q_{k,n} \in (0, 1)$ and $r_{k,n} \in \mathbb{N}$. The question is whether (2) holds for all probability generating functions (pgf's).

Looking for analogues of (*i*) and (*ii*) in Section 1, we see that (*i*) is automatically taken care of, but that (*ii*) does not apply: where for a distribution concentrated at $c \in \mathbb{R}_+$ we have

$$e^{-cs} = \lim_{n \to \infty} \left(\frac{1}{1 + \frac{c}{n}s} \right)^n,\tag{3}$$

the pgf z^k of a distribution concentrated at $k \in \mathbb{N}_+$ cannot be approximated by the right-hand side of (2) for the following reason. Since the function

$$\left(\frac{1-q}{1-qz}\right)^k = \left(\frac{1}{1-\lambda(1-z)}\right)^k,\tag{4}$$

with $\lambda = q/(1-q)$, is completely monotone (has alternating derivatives) in w := 1-z for $w \in \mathbb{R}_+$, and complete monotonicity is preserved under the taking of pointwise limits, the right-hand side of (2) is completely monotone. This means that complete monotonicity in 1-z is necessary for *P* to satisfy (2). Since, clearly, the function $z^k = (1-w)^k$ is not completely monotone in w on \mathbb{R}_+ , it cannot be obtained as (2).

We now turn to the sufficiency of the complete monotonicity in 1 - z. By Bernstein's theorem (Feller (1971)) a function *h* is completely monotone in *s* if and only if it can be represented as

$$h(s) = \int_{[0,\infty)} e^{-sx} dH(x),$$

where H is nondecreasing. It follows that a pgf P satisfying (2) must be of the form

$$P(z) = \int_{[0,\infty)} e^{-x(1-z)} dF(x),$$
(5)

where *F* is a df, i.e., *P* must be of the form $P(z) = \hat{F}(1 - z)$. But then (2) follows from (1), and we find that complete monotonicity of *P* as a function of 1 - z is not only necessary to have (2), but also sufficient. Summarizing we have the following result.

Theorem 1 A distribution $(p_k)_0^{\infty}$ on \mathbb{Z}_+ can be approximated by mixtures of negative binomial distributions if and only if (p_k) is a Poisson mixture, i.e., if and only if the pgf P of (p_k) satisfies (5).

3 Remarks and an example

A distribution on \mathbb{Z}_+ satisfying equation (5) is called a Poisson mixture; the corresponding random variable, N say, is of the form

$$N \stackrel{d}{=} N(X),$$

where $N(\cdot)$ is a unit Poisson process and X is an \mathbb{R}_+ -valued random variable independent of $N(\cdot)$. A detailed discussion of Poisson mixtures can be found in Puri and Goldie (1979). Between the moments of N and X we have the following relations.

$$EN = EX, \qquad EN(N-1) = EX^2,$$

and hence

$$\operatorname{var}(N) = EN + \operatorname{var}(X).$$

This means that *N* is not a Poisson mixture, and hence cannot be approximated by mixtures of negative binomial distributions if EN > var(N). The example in Section 2 demonstrates this; when P(N = k) = 1, then EN = k > 0 = var(N).

That mixtures of degenerate distributions on \mathbb{R}_+ have mixtures of Poisson distributions as their analogues on \mathbb{Z}_+ results from the fact that, in many respects, the Poisson distributions themselves are the analogues on \mathbb{Z}_+ of the degenerate distributions on \mathbb{R}_+ . The analogue of (3) is obtained by putting k = n, $\lambda = \frac{c}{n}$ in (4) and letting *n* tend to ∞ ; this yields the Poisson pgf exp(c(z-1)) as a limit. Similarly, the Gamma mixtures are taken into negative binomial mixtures by the following simple transformation.

$$\left(\frac{1}{1+c(1-z)}\right)^k = \frac{1}{k!} \int_0^\infty e^{-c(1-z)} x^k e^{-x} dx.$$

For another analogy between Poisson distributions and degenerate distributions we refer to Steutel and van Harn (1979).

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