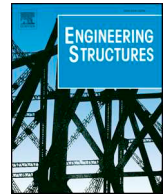




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# Evaluation of the variability contribution due to epistemic uncertainty on constitutive models in the definition of fragility curves of RC frames



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## ARTICLE INFO

### Keywords:

Reinforced Concrete buildings  
Modelling uncertainty  
Constitutive model  
Seismic fragilities  
ANOVA

## ABSTRACT

In the framework of uncertainty propagation in seismic analyses, most of the research efforts were devoted to quantifying and reducing uncertainties related to seismic input. However, also uncertainties associated to the definition of constitutive models must be taken into account, in order to have a reliable estimate of the total uncertainty in structural response. The present paper, by means of incremental dynamic analyses on reinforced concrete frames, evaluates the effect of the epistemic uncertainty for plastic-hinges hysteretic models selection. Eleven different hysteretic models, identified based on literature data, were used and seismic fragility curves were obtained for three different levels of maximum interstorey drift ratio. Finally, by means of analysis of variance techniques, the paper shows that the uncertainty associated to the hysteretic model definition has a magnitude similar to that due to record-to-record variability.

## 1. Introduction

Performance-Based Earthquake Engineering (PBEE) probabilistic procedures, aim at providing a framework for the systematic treatment of all the uncertainties involved in seismic engineering problems [1,2]. These uncertainties are mainly related to seismic actions, typically named record-to-record variability, and nonlinear structural behaviour in dynamic conditions. The latter uncertainties can be associated to different factors: (a) definition of material properties; (b) material hysteretic behaviour; (c) definition of backbone curves and cyclic behaviour of sections; (d) definition of numerical models (geometry uncertainty, load position, mass distribution, damping, description of damage mechanisms that can lead to collapse, etc.). These uncertainties are sometimes grouped under the label of modelling uncertainty [3].

Most of the vast literature on the topic concerns the quantification of the uncertainty in structural response due to record-to-record variability [4–14]. Whereas the level of knowledge on the effects of modelling uncertainty is much more limited. For instance, for many years, in seismic analyses, properties of materials, constitutive and geometric models, were assumed as deterministic [15–18]. Recently, some researchers investigated the effects of uncertainties in the definition of materials properties [19–21] or constitutive models [3,22], but the consequences of epistemic uncertainties on constitutive model selection (i.e. uncertainties related to the assumption of a specific constitutive model in a numerical simulation) were seldom studied.

This fact could lead to the false conclusion that effects of the latter are less important, or even negligible, than the effects of ground motion variability.

One of the reasons why modelling uncertainty was not investigated as much as record-to-record variability is probably its strong system-dependency (i.e. dependent on constructive technology) which makes not easy to obtain general results. Recently, the importance of considering modelling uncertainties, especially in the definition of collapse capacity has been acknowledged by researchers [23,24]. Vamvatsikos and Fragiadakis [22] evaluated the variability in both seismic demand and capacity, on a nine-storey steel moment resisting frame by means of Incremental Dynamic Analysis (IDA). They adopted moment-rotation relationships with non-deterministic backbones for plastic-hinges in beams, and assessed the seismic performance of their model for several combinations of parameters. The uncertain parameters of the backbones were: the yield moment, the post-yield hardening ratio, the end-of-hardening rotation, the slope of the softening branch, the residual moment capacity and the ultimate rotation. The same hysteretic rules were adopted in all models. Vamvatsikos and Fragiadakis have also confirmed that, because of the strongly non-linear nature of the problem, a model with median values of constitutive parameters, does not provide an estimate of the median seismic demand capacity. Dolsek [21] extended the standard IDA procedure by proposing to use a group of structural models, in addition to different ground-motion records, in order to estimate both modelling uncertainty and record-to-record

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<https://doi.org/10.1016/j.engstruct.2019.03.064>

Received 18 August 2018; Received in revised form 10 February 2019; Accepted 19 March 2019

Available online 26 March 2019

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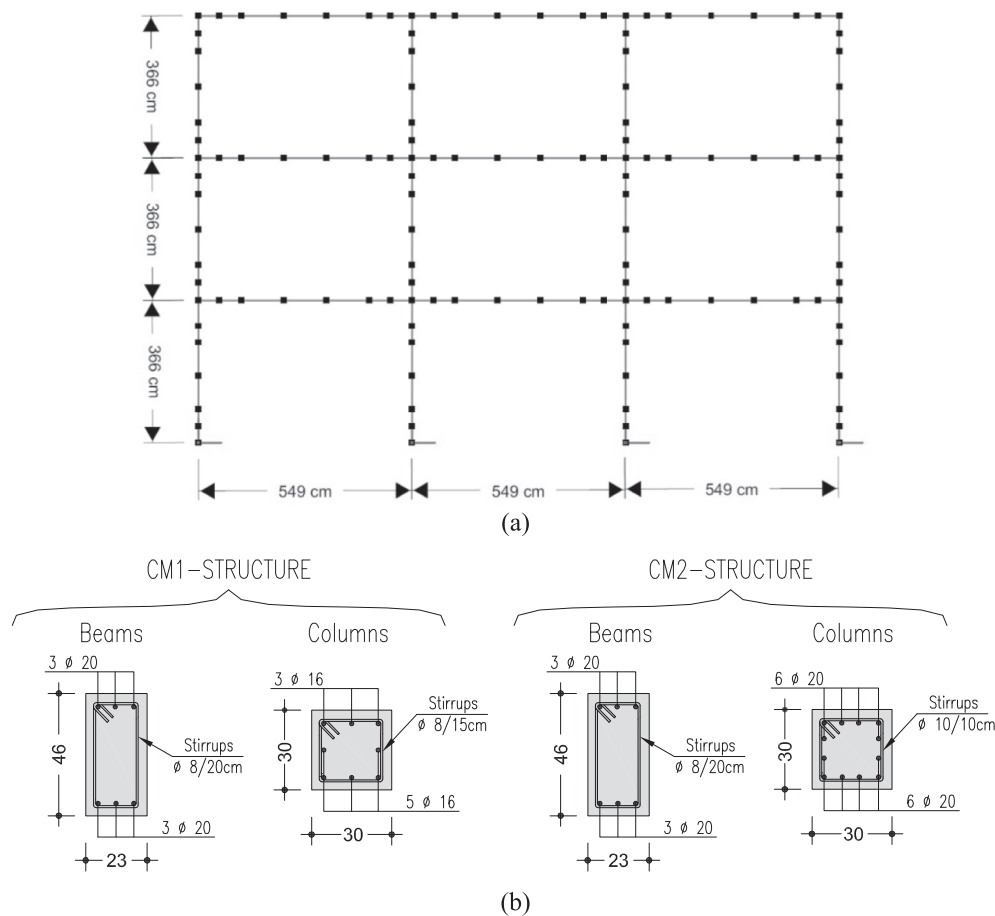


Fig. 1. Three-storey RC frame selected as case study. (a) FEM mesh. (b) Details of extremity transverse sections of the elements for the various designs considered.

variability. Considering a four-storey RC frame, Dolsek [21] has shown that the influence of modelling uncertainty has a significant influence on seismic collapse capacity. Borgonovo et al. [25] discussed importance measures in seismic probabilistic risk assessment and proposed a method to guide the choices of the modeller to detect critical factors in modelling, with the main goal of producing a suitable-to-purpose set of fragility curves obtained starting from a single fragility curve and considering the epistemic uncertainties affecting modelling activities. Celik and Ellingwood [26] by using probabilistic non-linear analyses, have proposed fragilities for Reinforced Concrete (RC) frames designed for gravity loads and concluded that record-to-record variability dominates the total uncertainty in the structural response of low-to-moderate seismic areas. Moreover, they stated that fragilities developed using mean values of structural parameters may be sufficient for earthquake damage and loss estimation in moderate seismic regions. This seems not completely in agreement with the findings in [3]. A similar conclusion is drawn by Kwon and Elnashai [20]. These authors, studying a three story ordinary moment resisting RC frame, indicated that the effect of randomness of material parameters had less importance than the effect of ground-motion variability. Recently, Ugurhan et al. [27], studied the uncertainty in the evaluation of the seismic collapse of modern RC frame structures, concluding that the probability of collapse is sensitive to modelling uncertainties.

The present paper discusses the effects of epistemic uncertainty on the choice of constitutive models for RC frame structures, while the uncertainty in model parameters is not addressed. In particular, in this work, IDA [28] is used in order to estimate the effect of uncertainties in the definition of hysteretic models for plastic hinges in beams and columns of RC frames. By analysing the results of the IDA analyses, we have obtained fragility curves for the maximum interstorey drift ratio,

in terms of both Peak Ground Acceleration (PGA) and spectral acceleration for the first natural period of the structures under consideration  $Sa(T_1)$ . Furthermore, by means of ANalysis Of VAriance (ANOVA) techniques, record-to-record variability and constitutive modelling variability are quantified, showing that, at large IDRs, the latter is comparable to the former.

## 2. The case study: three-storey RC moment resisting frame

### 2.1. Description of the structure and definition of collapse mechanisms

The structure used as case study is a three-storey three-bay RC frame (Fig. 1), with the same geometry adopted in [26]. This frame has 30 cm × 30 cm column cross-sections and 23 cm × 46 cm (B × H) beam cross-sections, the length of the bays is 549 cm and the interstorey height is 366 cm for all the three storeys. For the design against vertical loads, a distributed load of 48 kN/m was assumed on the beams. The design for vertical loads complies to Eurocode 2 prescriptions [29].

In order to reproduce the behaviour of different classes of structures, two alternative reinforcement bars (rebars) designs were considered. These designs can be associated to two different types of expected collapse mechanisms (named collapse mechanism CM1 and CM2 in the following). The CM1 design is characterized by weak column-strong beam failure, that can be representative of existing structures designed without the application of capacity design principles (in particular the strong column-weak beam hierarchy) [30]. The layout of reinforcement bars for this first structure is depicted in Fig. 1b. Beams have 3Ø20 bars at the top and at the bottom of the end sections. In the central portion of the beams there are 2Ø14 and 4Ø16 at the top and at the bottom, respectively. The columns have 3Ø16 rebars for each edge.

Based on this design a soft storey failure mechanism is expected.

The CM2 reinforcement design is fully compliant to Eurocode 8 prescriptions, thus including capacity design rules. Therefore, in this case a global mechanism of collapse is expected, with plastic hinges at the ends of beams. The beams of this frame contain the same rebars of those of CM1, while total of 12Ø20 are considered for column cross-sections (see Fig. 1b). For both reinforcement designs, the shear capacity of columns and beams was assumed higher than the flexural capacity, so to exclude brittle shear failures. Any other possible brittle failure mechanism (e.g. buckling and loss of anchorage of rebars, pull-out of the cover, joints failure) was excluded as well.

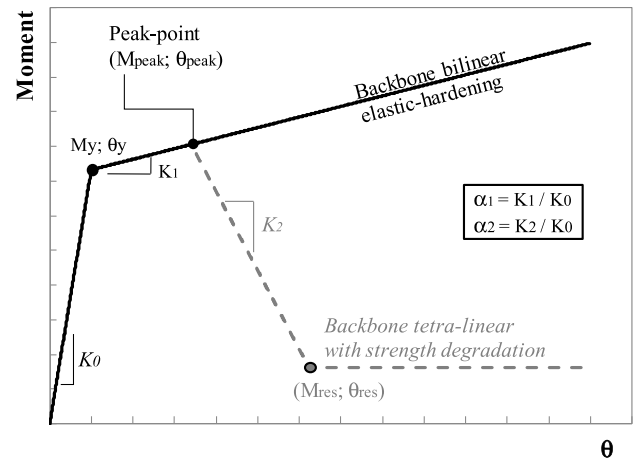
### 2.2. Finite element modelling

Numerical analyses were carried out using finite element models of the frames. The software OpenSees was employed to this aim [31]. Fig. 1a illustrates the main geometry of the structure considered. Beams and columns were modelled using elastic finite elements with lumped plastic hinges at their ends. At the ground level, columns were fully clamped. Masses corresponding to structural and non-structural dead loads and a fraction of live loads, according to Eurocode 8 criteria, were considered as equivalent distributed masses on the beams. A concrete compressive strength of 36 MPa was assumed as well as an elastic modulus of 30.46 GPa. The adopted steel yielding stress was 500 MPa and its elastic modulus 200 GPa. Both the frames considered had a natural period of  $T_1 = 1.09$  s.

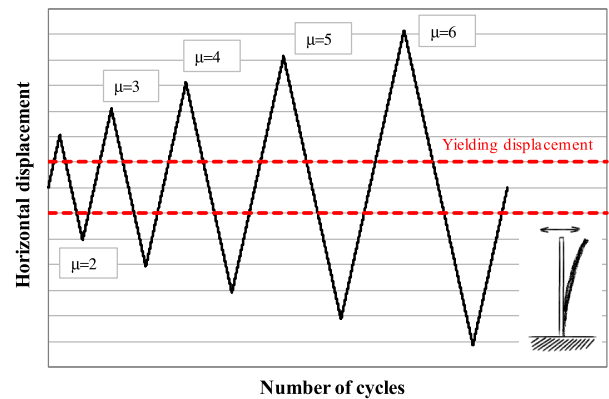
Nonlinearity was introduced in the FE models by considering cyclic nonlinear moment-rotation relationships for the plastic hinges. In order to investigate the effects of the uncertainty in the choice of the hysteretic model, eleven different rules were considered in the study, as discussed in Section 2.3. Furthermore, two types of moment rotation backbone curve were used (see Fig. 2a): (i) bilinear elastic-hardening and (ii) tetra-linear with strength degradation. The bilinear backbone and the first two branches of the tetra-linear model are defined by the same parameters, i.e. yielding moment ( $M_y$ ) and rotation ( $\theta_y$ ) and peak moment ( $M_{peak}$ ) and rotation ( $\theta_{peak}$ ). The tetra-linear model features a negative stiffness branch after the peak moment capacity. The coordinates of the points of these backbone curves were calculated as recommended by FEMA-273 [32], taking into account the effect of the axial load in static conditions in column cross-sections, and considering values recommended for elements with flexural dominated failure. The residual bending moment ( $M_{res}$ ) was assumed as 20% of  $M_{peak}$  ( $M_{peak}$  in Fig. 2a), as suggested in [32]. The rotation corresponding to the residual moment ( $\theta_{res}$ ) was considered as a fitting parameter.

The *Hysteretic* material model available in Opensees was adopted in the analyses. This model describes the moment-rotation relationship based on a tetra-linear backbone curve and five damage parameters related to cyclic degradation (i.e.  $P_x$ ,  $P_y$ , D1, D2,  $\beta$ ) [33]. The backbone curve of the model (see Fig. 2a) is defined as a function of the coordinates of the end points for the different linear branches, i.e. ( $M_y, \theta_y$ ), ( $M_{peak}, \theta_{peak}$ ) and ( $M_{res}, \theta_{res}$ ). By properly setting the values of these parameters it is possible to represent both the bilinear and tetra-linear backbones discussed above.

As for the damage parameters,  $P_x$  indicates the pinching factor for rotations (horizontal axis in a moment-rotation diagram) during re-loading, while  $P_y$  is the pinching factor for moments (vertical axis in a moment-rotation diagram). Unit values for  $P_x$  and  $P_y$  correspond to no pinching while smaller values will introduce the pinching effect. The parameters  $P_x$  and  $P_y$  are typically set together; for RC structural elements with pinching-like degradation mechanisms, the value of the ratio  $P_y/P_x$  is normally lower than 1.0 and spans from 0.7 (low pinching degradation) to 0.2 (severe pinching degradation). The parameters D1 and D2 are used to model cyclic strength degradation based on ductility and dissipated energy, respectively [33]. In particular, the strength reduction introduced by D1 is proportional to the maximum rotation reached during the loading cycles while the strength reduction



(a)



(b)

Fig. 2. Description of the parameters governing the backbone curves for the plastic hinges (a) and displacements history used in the calibration procedure (b).

associated to D2 is proportional to the hysteretic energy dissipated. The parameter  $\beta$  describes unloading stiffness degradation. Following a ductility-based approach [34], the unloading stiffness  $K_{UNL}$  is expressed by the following relationship:

$$K_{UNL} = K_0 \left( \frac{\theta_{max}}{\theta_y} \right)^\beta \tag{1}$$

where  $\theta_{max}$  indicates the maximum rotation in the plastic hinge under consideration. The calibration process used to define the values of the aforementioned parameters is discussed in Section 2.4.

### 2.3. Definition of observed degradation, damage index, damage parameters and dissipated energy ratio

#### 2.3.1. Definition of possible hysteretic models

The main scope of the present work is to evaluate and define, from a probabilistic point of view, the effects on seismic fragility curves of the epistemic uncertainty related to hysteretic model selection. Therefore, a set of cyclic constitutive models that represent possible alternatives must be defined. Constitutive models adopted in structural analyses are typically empirically defined based on direct observations or theoretically derived, therefore it is unlikely that a single constitutive model is perfectly representative of the behaviour of structural elements [35]. Even when detailed experimental data is available, the prediction of the seismic structural behaviour remains uncertain as many blind prediction contests have shown [36,37]. Moreover, if the structure under investigation is part of an existing building, uncertainty might be even

**Table 1**  
Constitutive models adopted in the study with the corresponding damage mechanisms.

Constitutive model definition	Constitutive model code	Backbone curve B or T (*)	Types of degrading mechanisms considered in the model						
			Reloading stiffness	Unloading stiffness	Cyclic strength	In-cycle strength	Pinching		
Clough-Johnston model	REF	B	●						
In-cycle strength degrading with DER = 65%	IN-CYC-65	T	●			●			
In-cycle strength degrading with DER = 85%	IN-CYC-85	T	●			●			
Unloading stiffness degrading with DER = 65%	UNL-65	B		●					
Unloading stiffness degrading with DER = 85%	UNL-85	B		●					
Cyclic strength degrading with DER = 65%	CYC-65	B			●				
Cyclic strength degrading with DER = 85%	CYC-85	B			●				
Pinching degrading with DER = 65%	PINCH-65	B						●	
Pinching degrading with DER = 85%	PINCH-85	T						●	
Mixed degrading with DER = 65%	MIX-65	T	●	●	●	●	●		
Mixed degrading with DER = 85%	MIX-85	T	●	●	●	●	●		

(\*) B: Bilinear elastic hardening; T: Tetra-linear with strength degradation.

**Table 2**

Values of the damage parameters used for the different constitutive models.

	$P_x$	$P_y$	D1	D2	$\beta$	$\alpha_2$	$\alpha_M$
<i>Damaging parameters for structures CM1</i>							
REF	1.0	1.0	0.0	0.0	0.0	+0.030	0.0
IN-CYC-65	1.0	1.0	0.0	0.0	0.0	-0.528	0.2
IN-CYC-85	1.0	1.0	0.0	0.0	0.0	-0.203	0.2
UNL-65	1.0	1.0	0.0	0.0	0.7	+0.030	0.0
UNL-85	1.0	1.0	0.0	0.0	0.45	+0.030	0.0
CYC-65	1.0	1.0	0.0	0.95	0.0	+0.030	0.0
CYC-85	1.0	1.0	0.0	0.26	0.0	+0.030	0.0
PINCH-65	0.93	0.20	0.0	0.0	0.0	+0.030	0.0
PINCH-85	0.95	0.68	0.0	0.0	0.0	+0.030	0.0
MIX-65	1.0	1.0	0.0	0.29	0.50	-0.120	0.2
MIX-85	1.0	1.0	0.0	0.13	0.25	-0.071	0.2
<i>Damaging parameters for structure CM2</i>							
REF	1.0	1.0	0.0	0.0	0.0	+0.0067	0.0
IN-CYC-65	1.0	1.0	0.0	0.0	0.0	-0.265	0.2
IN-CYC-85	1.0	1.0	0.0	0.0	0.0	-0.136	0.2
UNL-65	1.0	1.0	0.0	0.0	0.77	+0.0067	0.0
UNL-85	1.0	1.0	0.0	0.0	0.52	+0.0067	0.0
CYC-65	1.0	1.0	0.0	2.40	0.0	+0.0067	0.0
CYC-85	1.0	1.0	0.0	0.90	0.0	+0.0067	0.0
PINCH-65	0.93	0.19	0.0	0.0	0.0	+0.0067	0.0
PINCH-85	0.95	0.67	0.0	0.0	0.0	+0.0067	0.0
MIX-65	1.0	1.0	0.0	1.6	0.3	-0.0294	0.2
MIX-85	1.0	1.0	0.0	0.6	0.2	-0.0294	0.2

higher because of limited knowledge on material properties, geometry, detailing, etc. Therefore it is not possible to deterministically define constitutive models for structural members (or cross-sections). This matter is also more complicated when dealing with cyclic (seismic) loading because the expected hysteretic degradation mechanisms must be selected. This might be done using engineering judgment, but, the use of subjective expert opinions might lead to the tendency to include too many models which are often: (i) highly dependent, resulting in redundancy; or (ii) not plausible, resulting in a potentially significant over-estimation of model uncertainty [35,38].

The constitutive models and hysteretic rules adopted here were defined based on literature experimental outcomes. An analysis of the literature on experimental tests of RC elements [39–44], allowed to identify their main hysteretic degradation mechanisms. At least five mechanisms, relevant for the definition of the plastic hinge behaviour, were found: two involving stiffness (i.e. unloading stiffness degradation and reloading stiffness degradation), two affecting strength (i.e. cyclic strength degradation and in-cycle strength degradation) and a pinching type degradation (due to slip between concrete and rebars). The pinching-type degradation mechanism has distinctive features and can be considered as independent [45], whereas the other four mechanisms can act in different combinations. Kurtman [46] has identified and quantified the degrading behaviour of 196 RC cantilever elements, using the results of experimental cyclic tests extracted from the PEER Structural Performance Database [39] and from [40]. Furthermore, Kurtman found that the reloading stiffness degradation mechanism was present in almost all the considered cases.

The literature suggests that, among the multilinear models not requiring hysteretic damage parameters (i.e. elastic-plastic model, elastic-hardening bilinear model, peak oriented model and Clough and Johnston model) the Clough and Johnston model [47] is the most suitable to represent the flexural behaviour of RC structural elements with no or very-limited cyclic degradation. Therefore, Clough and Johnston hysteretic behaviour considering only reloading stiffness degradation was assumed as reference model (labelled as REF code in Table 1) for the calibration of all the other hysteretic models. The Clough and Johnston model was obtained by setting the parameters  $P_x = P_y = 1.0$  and  $D1 = D2 = \beta = 0.0$ . Based on the analysis of the literature the eleven different hysteretic models reported in Table 1

were considered in this study. For each of them Table 1 shows the degradation mechanisms taken into account and the type of backbone curve used.

The UNL (unloading) class of models represent the behaviour of elements with reloading and unloading stiffness degrading mechanisms. This class of models were obtained by calibrating the parameter  $\beta$  with

the other damaging parameters set as for the REF case. The CYC (cyclic) class has been added with the aim to consider cyclic strength and reloading stiffness degradation mechanisms, identifying a proper value of the parameter D2. The IN-CYC (in-cycle) models reproduce the hysteretic behaviour of degrading hinges for in-cycle strength and reloading stiffness decay mechanisms. This class of models were obtained

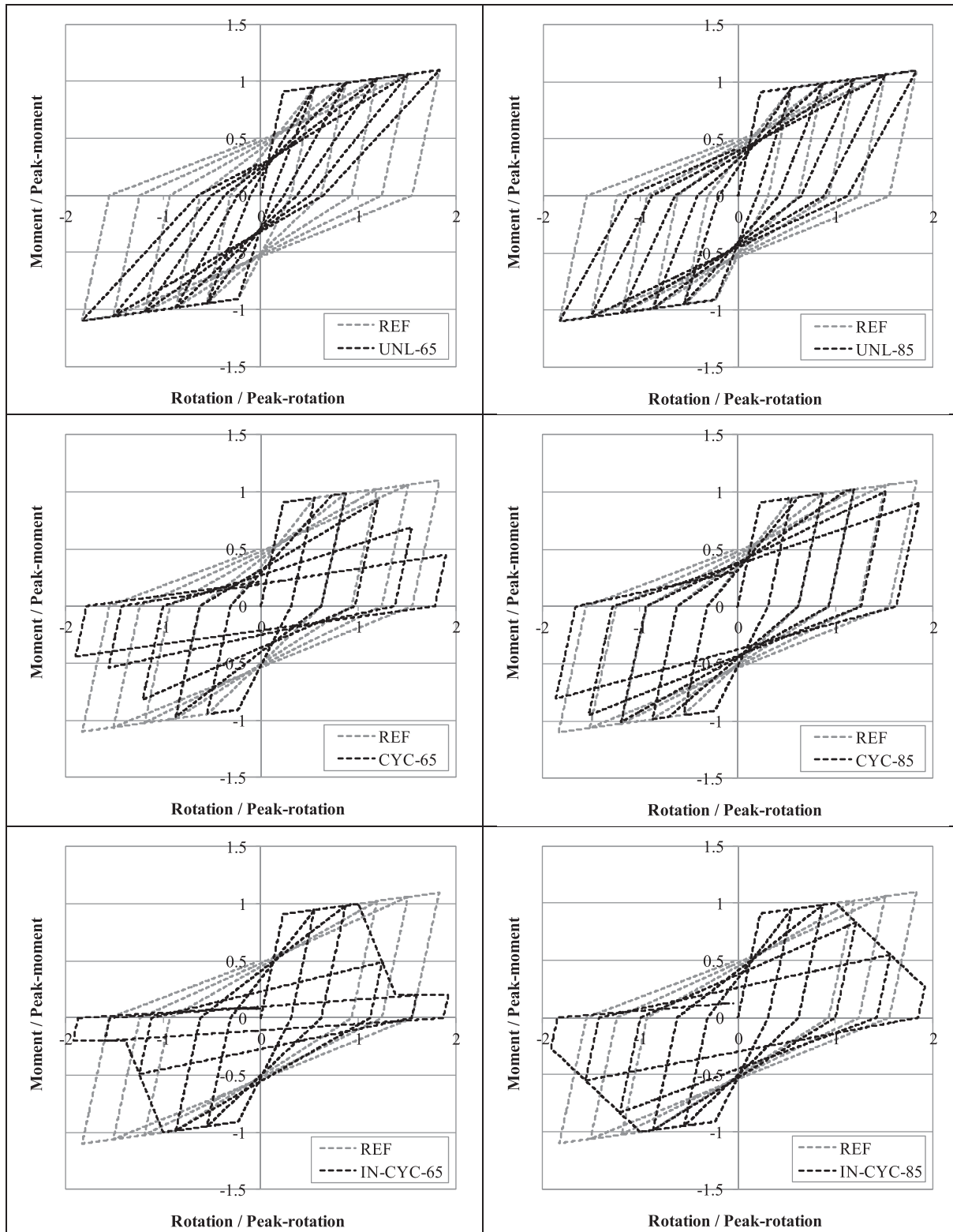


Fig. 3. Hysteretic moment-rotation relationships used for the various constitutive models for the frame CM1 (a) and CM2 (b). For a better comparison moments and rotations were normalized to their peak values.

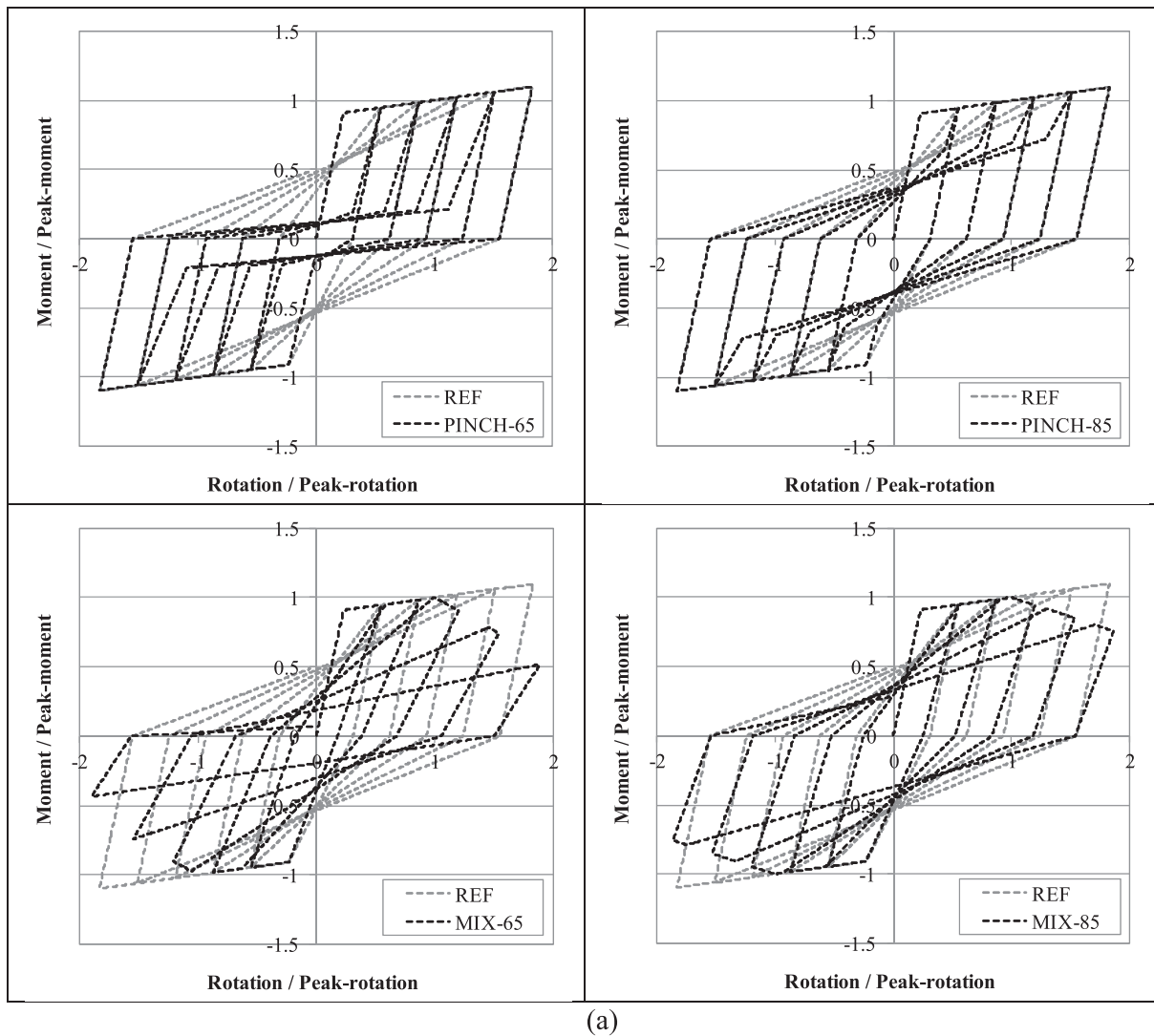


Fig. 3. (continued)

by using backbone curves with strength degradation (tetra-linear). The rotation  $\theta_{res}$  was considered as unknown. The pinching-type degradation mechanisms are represented by the class PINCH. In this case the calibration process identified the proper value of the ratio  $P_Y/P_X$ . Finally, a sixth class of models was adopted, in order to consider the contemporary presence of unloading-reloading stiffness degradation and cycle and in-cycle strength degradation (MIX). For the calibration process of this class of models  $\theta_{res}$  was considered as unknown (see Fig. 2a), having fixed the values of  $\beta$  and D2 on the basis of the experimental evidence collected in [46] and assuming as backbone the curve with strength degradation.

### 2.3.2. Dissipated energy measure

In order to compare the different hysteretic behaviours under consideration a proper damage measure needs to be defined. Several local damage indices are available in literature [48–52]; some of them depend on the ratio of the total dissipated energy over the total energy available, other parameters depend either on plastic rotation or curvature demand and capacity [53]. Anyway, they seem not fully suitable for the aim of the paper, since they express a ratio between demand and capacity evaluated on a specific constitutive model and therefore do not allow to compare different models. For the aim of the present work a parameter based on the dissipated hysteretic energy was defined. In particular, assuming the Clough and Johnston model as reference, a

cumulated Dissipated Energy Ratio (DER) was defined, for a general hysteretic model, as:

$$DER_n [\%] = \frac{\sum_{i=1}^n E_i}{\sum_{i=1}^n E_i^{CJ}} \cdot 100 [\%] \quad (2)$$

where  $E_i$  and  $E_i^{CJ}$  represent, the hysteretic energies dissipated under a prescribed loading sequence up to the  $n$ -th loading cycle, for a general constitutive model and for the Clough and Johnston constitutive model, respectively. Based on experimental data collected in [46], DER values can range from 80 to 90% for structural elements with low to moderate hysteretic damage to a minimum of 55–60% for members with severe hysteretic damage.

### 2.4. Calibration of the hysteretic behaviour of plastic hinges

Moment-rotation backbone curves presented in Sections 2.2 define the behaviour of the plastic hinges under increasing monotonic loading, but they are not sufficient to describe the hysteretic behaviour of RC elements under cyclic reverse loading. To this aim it is necessary to define values for the cyclic damage parameters (see Section 2.3). These were calculated by means of an inverse analysis procedure, based on the results of the wide statistical analysis of the results from [39,40]. This procedure considers a simple cantilever column model, comprising

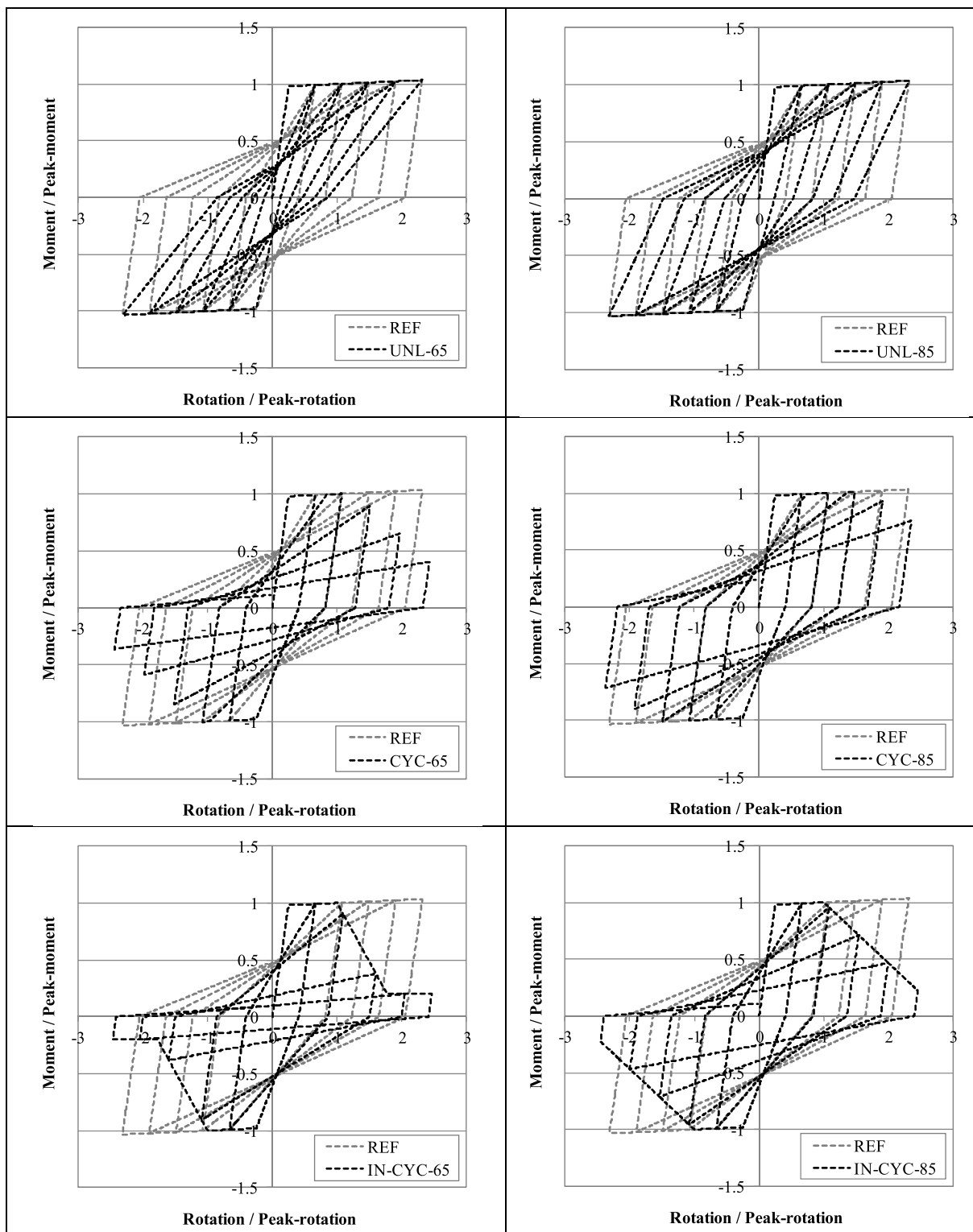
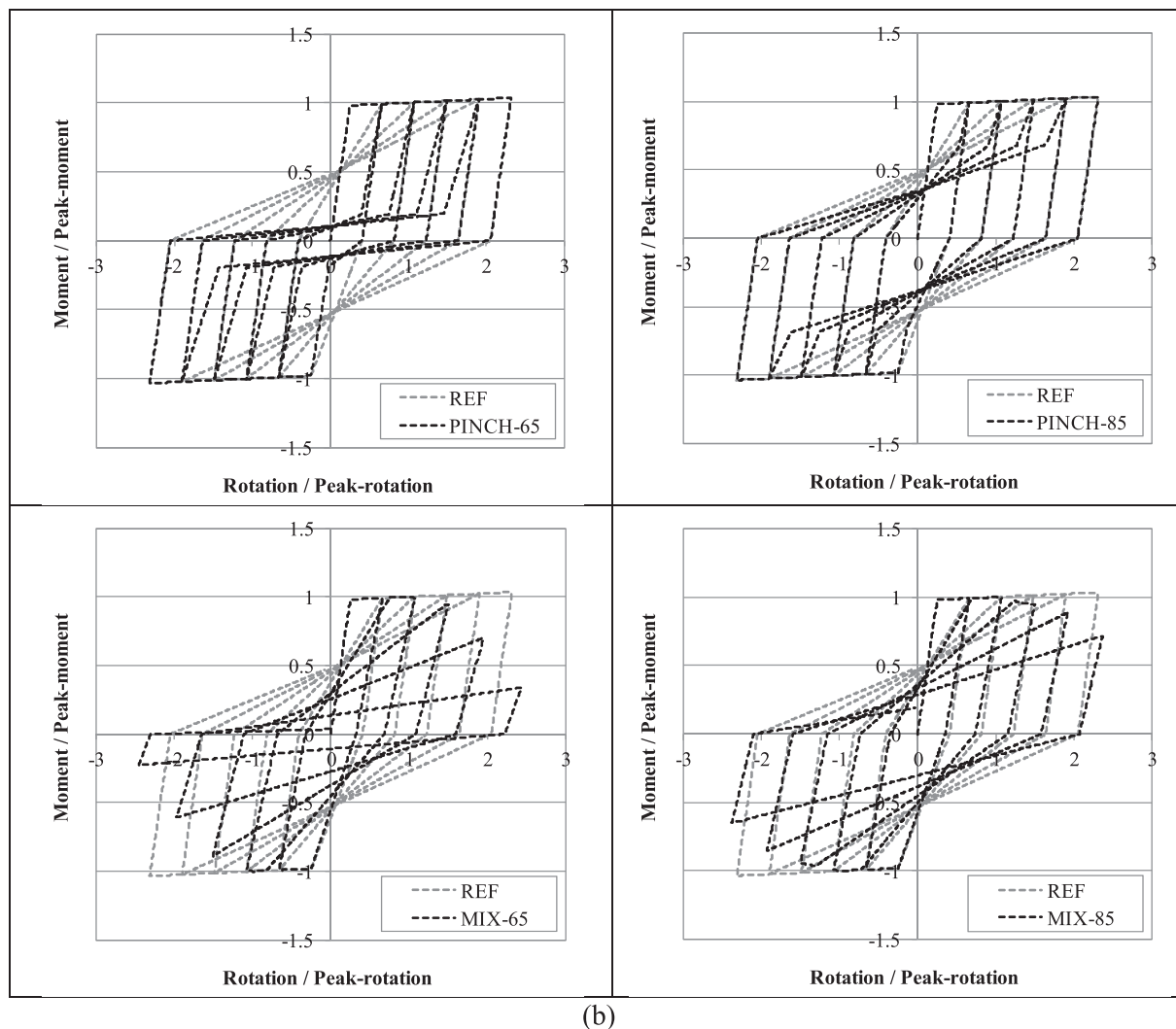


Fig. 3. (continued)

an elastic element and a nonlinear hinge at the base. The hinge has a backbone curve derived from the properties of either the cross-sections of columns, model CM1, or the end sections of beams, model CM2. Cyclic horizontal displacements with increasing amplitude were applied at the top of the cantilever column, for a total of five complete cycles. The cyclic displacement history was set in order to achieve displacement ductility demands of 2, 3, 4, 5 and 6, respectively (see Fig. 2b). The calibration procedure starts from the calculation of the

energy dissipated by the cantilever adopting the Clough-Johnston model (REF in Table 1). The parameters of the other ten models were then defined assuming considering target values for the DER: either 0.65 or 0.85. Table 2 reports the values of the damage parameters identified, and then used in the dynamic analyses, for each constitutive model. Table 2 also reports the values of the parameters  $\alpha_2$  and  $\alpha_M$  (see Fig. 2a where:  $\alpha_2 = K_2/K_0$  and  $\alpha_M = M_{res}/M_{peak}$  respectively) derived by the values of moments and rotations.



(b)

Fig. 3. (continued)

The moment-rotation numerical responses obtained for the various models for the cyclic quasi-static history of displacements are depicted in Fig. 3, overlapped to the response of REF model. For the ease of comparison, moments and rotations are normalized to the values corresponding to the peak-point of the backbone curve. Comparing the hysteretic behaviours depicted in Fig. 3 with the experimental results available in the literature, it is possible to notice that the group of plastic hinges with a DER value of 65% fit very well with the expected behaviour of members with poor detailing and confinement. On the other hand, the hysteretic behaviours calibrated to provide DER = 85% are more representative of elements complying with modern code-prescriptions. The effects of the single degradation mechanisms are clearly evidenced by the hysteretic behaviour of the classes UNL, CYC, IN-CYC and PINCH. The outcomes reported in the following sections will show which degrading mechanisms are able to produce the worst scenario for the investigated models. Based on these findings, it is possible to define a hierarchy of the most severe mechanisms potentially addressing future studies.

### 3. Seismic input and dynamic analyses

The dynamic performances of the structures were evaluated by IDAs. These analyses allow to evaluate the response variability caused by constitutive behaviour uncertainty for different levels of

displacement demand. The set of thirty ground-motion records selected in [22] were used in the present work. They represent a seismological scenario consisting of a magnitude 6.5–6.9 earthquake at a distance of 15–33 km (fault rupture distance). No record contains near-source effects (i.e. directivity effects). The main characteristics of the records are reported in Table 3, while Fig. 4 shows their elastic pseudo-acceleration response spectra for 5% damping ratio. This set of ground motion records was adopted in order to allow an easy comparison with other literature results. However, it should be noticed that ground-motion records selected according to a seismological scenario [54] (i.e. based on earthquake magnitude and distance only) might lead to an over-estimate of record-to-record variability when using IDA analyses, mainly because: i) the reference seismological scenario used for the selection (which can be obtained from hazard disaggregation) might not be representative for all the ground-motion IM values used in IDA; ii) information on spectral shape is not used for the selection [55], therefore the IDA procedure might not be robust against scaling. Other ground-motion selection criteria, such as those proposed by Baker [56] and by Bradley [57], could lead to a more accurate estimate of record to record variability.

In this study two different ground motion intensity measures (IM) were considered: the peak ground acceleration (PGA) and the 5%-damped spectral acceleration at the first natural period of vibration,  $S_a(T_1, 5\%)$ . The PGA is clearly a much less efficient IM than  $S_a(T_1)$  and

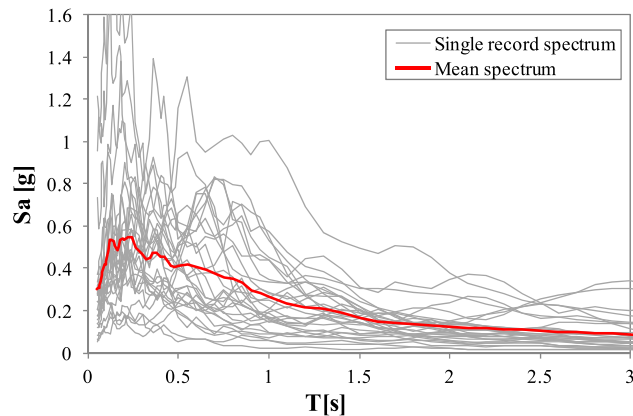


**Table 3**  
Set of ground motions adopted in the IDA procedure.

No.	Event name	Station	Component [°]	Soil	M*	R** [km]	PGA [g]
1	Loma Prieta, 1989	Agnews State Hospital	090	C	6.9	28.2	0.159
2	Northridge, 1994	LA, Baldwin Hills	090	B	6.7	31.3	0.239
3	Imperial Valley, 1979	Compuertas	285	C	6.5	32.6	0.147
4	Imperial Valley, 1979	Plaster City	135	C	6.5	31.7	0.057
5	Loma Prieta, 1989	Hollister Diff. Array	255	D	6.9	25.8	0.279
6	San Fernando, 1971	LA, Hollywood Stor. Lot	180	D	6.6	21.2	0.174
7	Loma Prieta, 1989	Anderson Dam Downstrm	270	B	6.9	21.4	0.244
8	Loma Prieta, 1989	Coyote Lake Dam Downstrm	285	B	6.9	22.3	0.179
9	Imperial Valley, 1979	El Centro Array #12	140	C	6.5	18.2	0.143
10	Imperial Valley, 1979	Cucapah	085	C	6.5	23.6	0.309
11	Northridge, 1994	LA, Hollywood Storage FF	360	C	6.7	25.5	0.358
12	Loma Prieta, 1989	Sunnyvale Colton Ave	270	C	6.9	28.8	0.207
13	Loma Prieta, 1989	Anderson Dam Downstrm	360	B	6.9	21.4	0.240
14	Imperial Valley, 1979	Chihuahua	012	C	6.5	28.7	0.270
15	Imperial Valley, 1979	El Centro Array #13	140	C	6.5	21.9	0.117
16	Imperial Valley, 1979	Westmoreland Fire Station	090	C	6.5	15.1	0.074
17	Loma Prieta, 1989	Hollister South & Pine	000	D	6.9	28.8	0.371
18	Loma Prieta, 1989	Sunnyvale Colton Ave	360	C	6.9	28.8	0.209
19	Superstition Hills, 1987	Wildlife Liquefaction Array	090	C	6.7	24.4	0.180
20	Imperial Valley, 1979	Chihuahua	282	C	6.5	28.7	0.254
21	Imperial Valley, 1979	El Centro Array #13	230	C	6.5	21.9	0.139
22	Imperial Valley, 1979	Westmoreland Fire Station	180	C	6.5	15.1	0.110
23	Loma Prieta, 1989	Halls Valley	090	C	6.9	31.6	0.103
24	Loma Prieta, 1989	WAHO	000	D	6.9	16.9	0.370
25	Superstition Hills, 1987	Wildlife Liquefaction Array	360	C	6.7	24.4	0.200
26	Imperial Valley, 1979	Compuertas	015	C	6.5	32.6	0.186
27	Imperial Valley, 1979	Plaster City	045	C	6.5	31.7	0.042
28	Loma Prieta, 1989	Hollister Diff. Array	165	D	6.9	25.8	0.269
29	San Fernando, 1971	LA, Hollywood Stor. Lot	090	C	6.6	21.2	0.210
30	Loma Prieta, 1989	WAHO	090	D	6.9	16.9	0.638

\* Moment magnitude.

\*\* Closest distance to fault rupture.



**Fig. 4.** Elastic pseudo-acceleration response spectra of the ground-motion records used in the present study.

therefore will lead to an higher record-to-record variability, nevertheless it was adopted here because it is still often adopted in the literature for the definition of fragility models.

The maximum interstorey drift ratio (IDR) over all the storeys was selected as engineering demand parameter (EDP), in order to define the response of the models. The IDAs were performed by scaling each record starting from a value of IM equal to 0.05 g until structural collapse, which was identified as either the achievement of the collapse IDR value, set equal to 5%, or the onset of dynamic instability. In this latter case, as in [22], flat IDA graph was considered until the collapse IDR (i.e. 5%). P-Δ effects were included in the analyses. A total of 1320 IDA curves were obtained and then analysed. In all the analyses maximum rotations resulted smaller than the ultimate rotation values given by the procedure described by Fajfar et al. [58].

#### 4. IDA analyses results

The present Section analyses the effects of each parameter and discusses the main outcomes of the uncertainty analysis. In the following, the individual IDA curves for the various ground-motion record are presented for the different types of structural models, together with their 16%, 50%, 84% percentile curves. Fig. 5a reports the results of the IDAs performed on the model CM1 adopting the REF constitutive model by considering PGA as IM, while Fig. 5b shows results in terms of  $S_a(T_1, 5\%)$ . As expected [28,59] the spectral acceleration is more efficient (i.e. provides values less scattered) than PGA for every level of interstorey drift considered. The dispersion in the values obtained using PGA is not negligible also in the quasi-elastic state of the frame. Fig. 5c and d show, for the model CM2 with the REF cyclic behaviour, the PGA and  $S_a(T_1, 5\%)$  IDA curves, respectively. Comparing the response of structure CM1 (Fig. 5b) and CM2 (Fig. 5d) one can observe that for the REF model, the 16%, 50% and 84% percentile curves have similar values in the IDR range considered.

In order to compare with more detail the variations in the response induced by the different constitutive models, Fig. 6a–d show IDA curves obtained for the frame CM2 for the constitutive models IN-CYC-65, UNL-65, PINCH-65 and MIX-85. In this case there is a relevant difference on the median collapse spectral acceleration (i.e. for IDR = 0.05). It is equal to 1.1 g for the IN-CYC-65, to 0.75 g for UNL-65 and to 0.7 g for the PINCH-65 model. It is worth noting that these three models correspond to the same level of DER = 65%. Furthermore, the model MIX-85, that combines various deterioration mechanisms, provides a lower median collapse spectral acceleration, equal to 0.68 g, even if it has a higher dissipation capacity (DER = 85%) than the other models (DER = 65%). This is not the only case among those considered in the study, therefore we can conclude with the adopted damage index (DER) it might not be possible to define a-priori which is the most

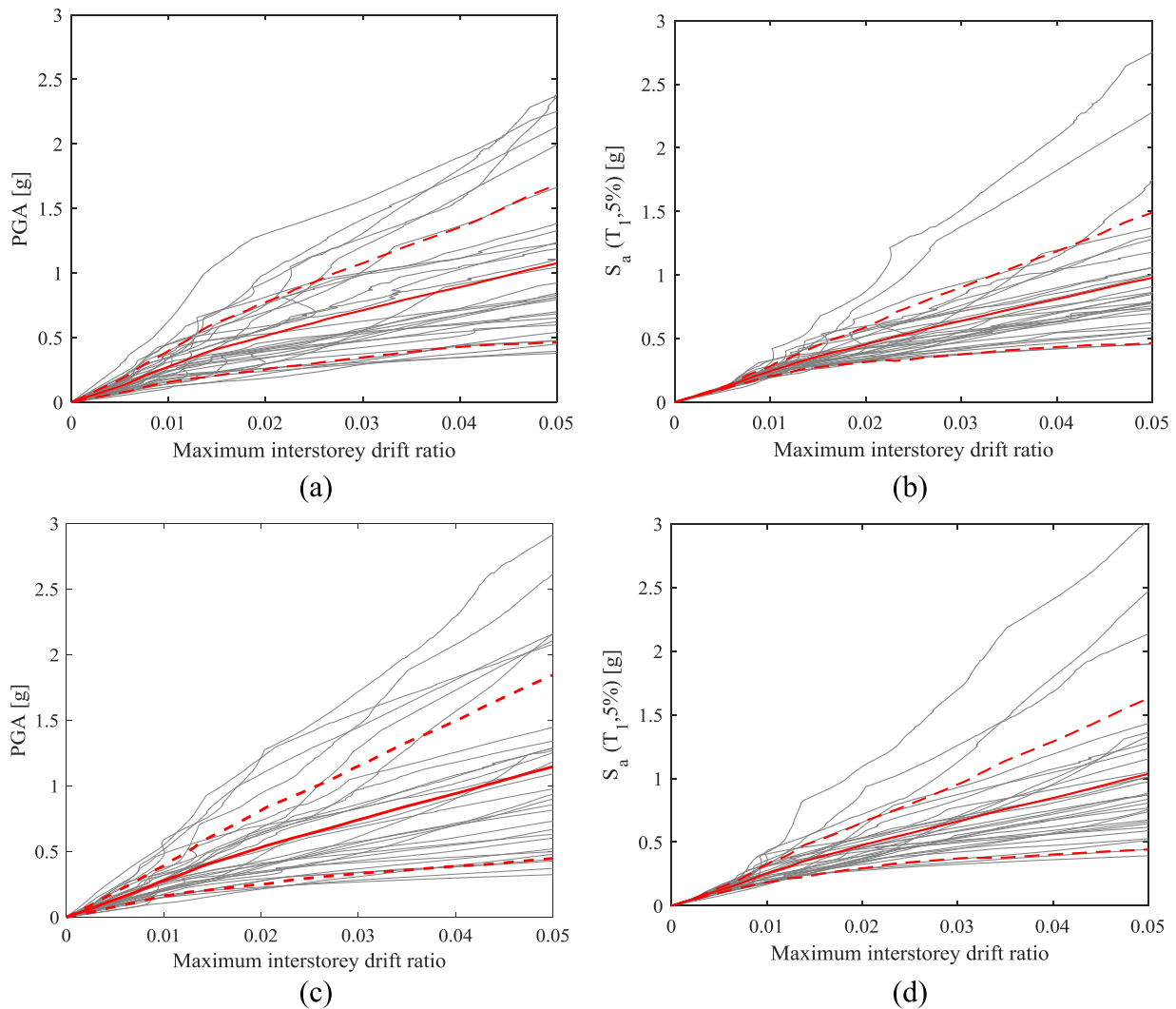


Fig. 5. IDA curves for single records (grey) and 16%, 50% and 84% percentiles (red) obtained from the REF constitutive model: (a) frame CM1 and IM = PGA, (b) frame CM1 and IM =  $S_a(T_1,5\%)$ , (c) frame CM2 and IM = PGA and (d) frame CM2 and IM =  $S_a(T_1,5\%)$ .

unfavourable constitutive model for the structures under consideration. It is important to highlight that the MIX model, produces IDA curves with a markedly irregular trend, most likely because the high complexity of the model leads to a strong sensitivity of the interstorey drift ratio to the ground-acceleration time variations, which cannot be described by a scalar IM.

In order to have a comprehensive overview of the main outcomes connected to the adoption of the different constitutive models Fig. 7 shows the median IDA curves for the different constitutive behaviours together with their mean curve, for the different CM. In general curves IDR vs. PGA result more scattered than IDR vs.  $S_a(T_1,5\%)$ . The structures have collapse spectral acceleration  $S_a(T_1,5\%)$  ranging between 0.35 g and 1.0 g. It is worth noting that the models IN-CYC-65, IN-CYC-85 and MIX-65 in Fig. 7a and b, having a softening behaviour with in-cyclic strength degradation, produces for structure CM1 very flat curves after IDR around 2% and until collapse (i.e. IDR = 5%). This trends occur because of the arising of a soft-storey mechanism at lowest floor and then the acceleration level activating the maximum capacity of the base plastic hinges is practically the collapse acceleration. Anyway, by comparing the various median IDA curves, is not possible to draw general considerations on the constitutive models producing the dynamic response reaching the highest/lowest structural capacities.

### 5. Fragility functions

This section presents fragility curves associated to each constitutive model, for the two different classes of structures (CM1 and CM2) considered in the paper. Following a consolidated method, the fragility function for a general IDR value is assumed to be a lognormal cumulative distribution function:

$$P(IDR = idr|im) = \Phi\left(\frac{\ln(im) - \mu}{\sigma}\right) \tag{3}$$

where  $im$  is either  $S_a(T_1)$  or PGA and  $\Phi(\cdot)$  is the standard normal cumulative distribution function. The values of the two parameters  $\mu$  and  $\sigma$ , defining the fragility function, can be estimated using different techniques [60]. In the present paper the maximum likelihood estimation method was used. Considering the general IDR limit (either 1%, 3% or 5% in the present paper) and defining a vector  $Y$  containing the  $N$  values if  $im$  corresponding to this limit, the likelihood function for the observations in  $Y$  can be defined as:

$$L(\mu, \sigma|Y) = \prod_{i=1}^N \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(\frac{-(\ln(y_i) - \mu)^2}{2\sigma^2}\right) \tag{4}$$

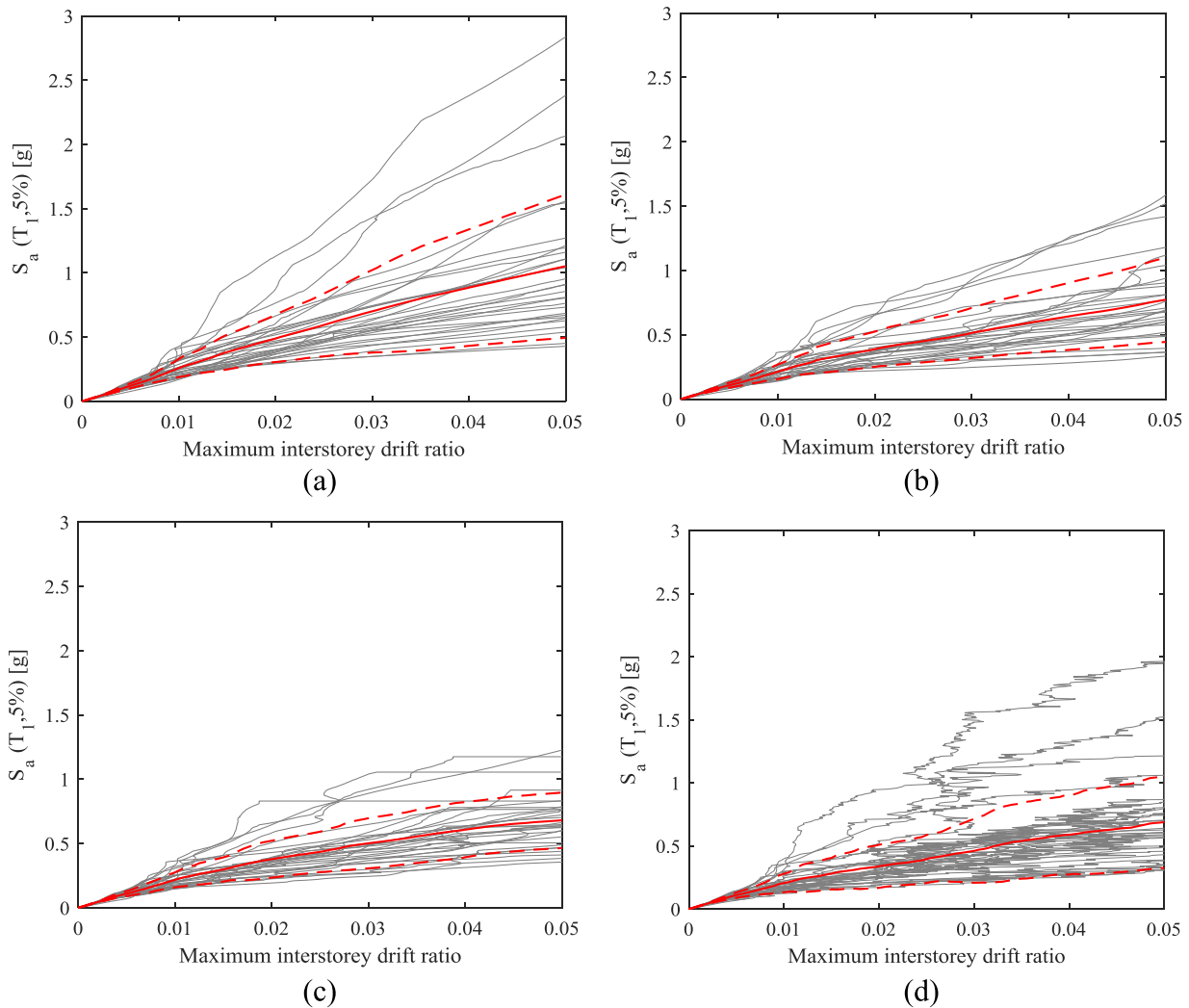


Fig. 6. IDA curves for single records (grey) and 16%, 50% and 84% percentiles (red) for the frame CM2 and the hysteretic models (a) IN-CYC-65, (b) UNL-65, (c) PINCH-65 and (d) MIX-85.

The estimates of the fragility model parameters ( $\mu$  and  $\sigma$ ) are those maximizing the likelihood function.

Fragility curves for the three IDR thresholds under consideration (representing slight damage, severe damage and collapse, respectively) were estimated both in terms of PGA and spectral acceleration  $S_a(T_1)$ . The median values of the fragility models,  $e^\mu$ , are reported in Tables 4 and 5 for the frames CM1 and CM2, respectively, and their logarithmic standard deviations (i.e. record to record variability) are reported in Tables 6 and 7. As an example, Fig. 8 shows some fragilities obtained for the structure CM2. In particular, two curves are plotted for each constitutive model: (i) the empirical cumulative distribution of structural capacity and (ii) the associated best fitting lognormal cumulative distribution. Fig. 8a shows results for IDR = 1% while Fig. 8b for IDR = 5%. Fig. 8, consistently with the results reported in Fig. 7, highlights the effects of the adoption of different constitutive models, both in terms of central value (i.e. median value of the fragility function) and in terms of variability. Considering the frame CM2 and IDR = 5% (Fig. 8b) the median values of  $S_a(T_1)$ , i.e.  $e^\mu$ , span between 0.336 g and 1.052 g, confirming the significant variability introduced by hysteretic models. This effect is relevant also for lower values of interstorey drift. For example, for frame CM2 and IDR = 1% (Fig. 8a) the median value of the curves ranges from 0.175 g to 0.259 g. It is worth noticing that the introduction of the in-cyclic strength degradation mechanism alone does not produce significant negative effects on the dynamic response of the structures. On the other hand, cyclic

strength degradation mechanisms affect structural capacity. It should also be noticed that, as expected, record to record variability ( $\sigma$ ) depends on the constitutive model adopted. Considering  $S_a(T_1)$  as IM, the frame CM2 and IDR = 1% the values of  $\sigma$  range from 0.246 (PINCH-65) to 0.348 (IN-CYC-65), while for IDR = 5% from 0.305 (PINCH-65) to 0.502 (IN-CYC-65).

Then, it might be of interest to define global fragility curves for the structures, by considering all the structural capacity values obtained from the various constitutive models. These global empirical fragilities represent a more general set of curves. Fig. 9 shows the global empirical fragility curves obtained for the various structures in terms of spectral acceleration. Fig. 9a shows the results for IDR = 1% (slight damage condition) while Fig. 9b and c the fragilities for IDR = 3% (i.e. severe damage condition) and IDR = 5% (i.e. collapse condition), respectively. The median values ( $e^\mu$ ) and logarithmic standard deviation ( $\sigma$ ) of the fitting lognormal distributions are summarized in Table 8, for both the intensity measures considered in the paper, i.e. PGA and  $S_a(T_1)$ . The median values for the two structures is similar for IDR = 1%. While the logarithmic standard deviation ranges from 0.18 g (CM1) to 0.31 g (CM2). The next Section will show that for this IDR limit most of the standard deviation is due to record-to-record variability. Considering an IDR threshold of 3% the structures CM1 and CM2 show again a comparable median capacity. If compared with the data for IDR = 1%, total standard deviation is now significantly larger but similar for the two structures. Finally, for IDR = 5%, the structure CM2 designed by

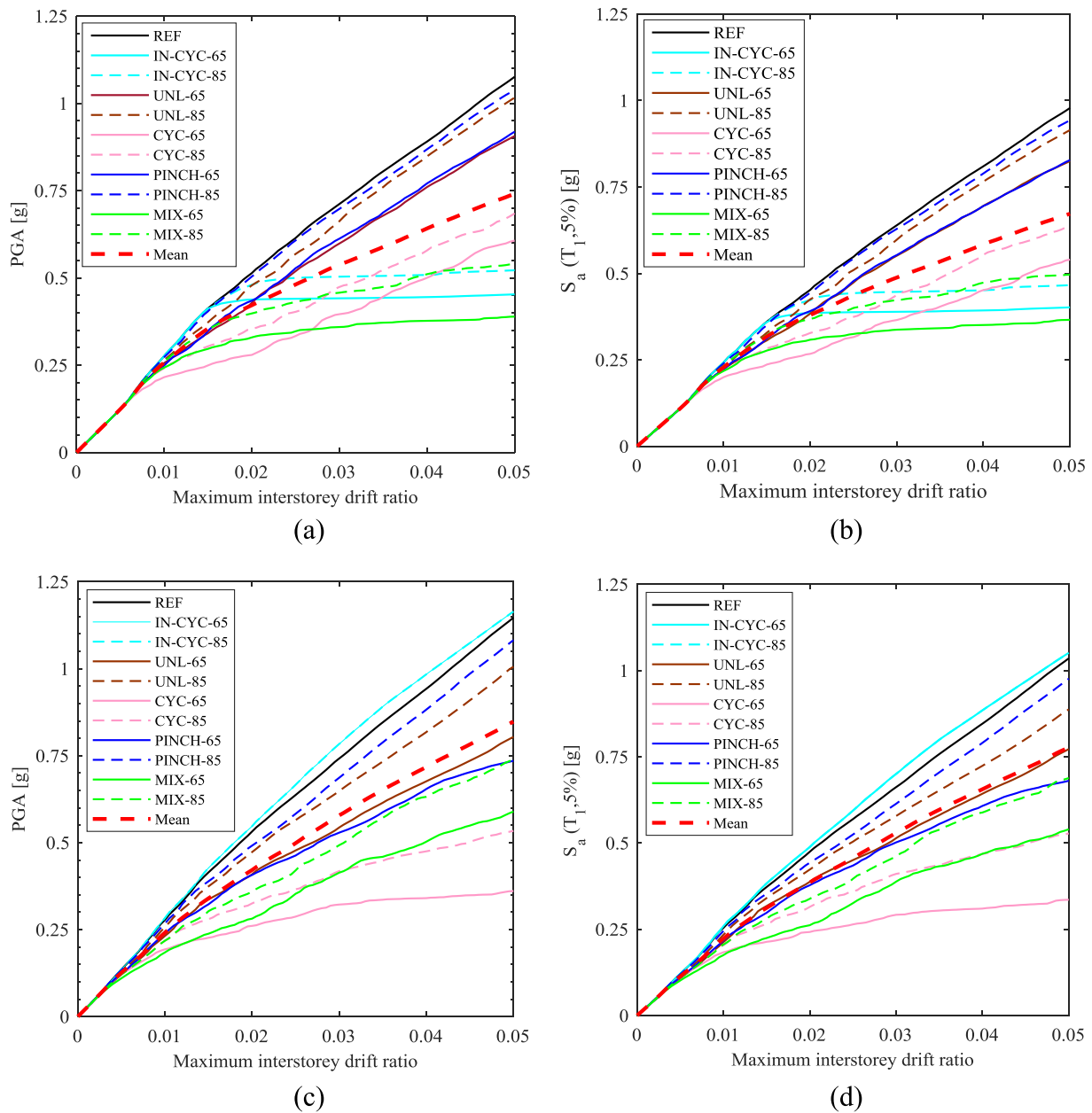


Fig. 7. Comparison of the median IDA curves (50% percentile) for the 11 constitutive models considered for (a) frame CM1 and IM = PGA, (b) frame CM1 and IM =  $S_a(T_1, 5\%)$ , (c) frame CM2 and IM = PGA and (d) frame CM2 and IM =  $S_a(T_1, 5\%)$ .

Table 4  
Median values ( $e^\mu$ ) of the fragility curves fitted using the different constitutive models for the frame CM1.

	IDR = 1%		IDR = 3%		IDR = 5%	
	PGA [g]	$S_a(T_1, 5\%)$ [g]	PGA [g]	$S_a(T_1, 5\%)$ [g]	PGA [g]	$S_a(T_1, 5\%)$ [g]
REF	0.276	0.243	0.712	0.639	1.076	0.978
IN-CYC-65	0.276	0.243	0.441	0.389	0.453	0.402
IN-CYC-85	0.276	0.243	0.503	0.447	0.522	0.466
UNL-65	0.256	0.225	0.598	0.551	0.906	0.825
UNL-85	0.265	0.232	0.662	0.598	1.016	0.915
CYC-65	0.215	0.200	0.396	0.365	0.607	0.540
CYC-85	0.241	0.220	0.474	0.437	0.685	0.639
PINCH-65	0.249	0.222	0.612	0.553	0.919	0.828
PINCH-85	0.273	0.240	0.698	0.629	1.038	0.943
MIX-65	0.243	0.218	0.359	0.337	0.389	0.366
MIX-85	0.261	0.233	0.457	0.423	0.539	0.496

**Table 5**  
Median values ( $e^\mu$ ) of the fragility curves fitted using the different constitutive models for the frame CM2.

	IDR = 1%		IDR = 3%		IDR = 5%	
	PGA [g]	Sa( $T_1,5\%$ ) [g]	PGA [g]	Sa( $T_1,5\%$ ) [g]	PGA [g]	Sa( $T_1,5\%$ ) [g]
REF	0.278	0.254	0.742	0.662	1.147	1.036
IN-CYC-65	0.283	0.259	0.784	0.700	1.164	1.052
IN-CYC-85	0.283	0.259	0.784	0.700	1.164	1.052
UNL-65	0.232	0.213	0.544	0.514	0.804	0.772
UNL-85	0.257	0.234	0.648	0.578	1.006	0.888
CYC-65	0.193	0.183	0.322	0.292	0.361	0.336
CYC-85	0.217	0.204	0.418	0.411	0.534	0.528
PINCH-65	0.239	0.217	0.529	0.502	0.735	0.681
PINCH-85	0.267	0.245	0.687	0.614	1.082	0.977
MIX-65	0.184	0.175	0.414	0.388	0.589	0.540
MIX-85	0.218	0.206	0.493	0.461	0.740	0.689

**Table 6**  
Logarithmic standard deviations ( $\sigma$ ) of the fragility curves fitted using the different constitutive models for the frame CM1.

	IDR = 1%		IDR = 3%		IDR = 5%	
	PGA [g]	Sa( $T_1,5\%$ ) [g]	PGA [g]	Sa( $T_1,5\%$ ) [g]	PGA [g]	Sa( $T_1,5\%$ ) [g]
REF	0.429	0.145	0.492	0.335	0.538	0.422
UNL-65	0.429	0.145	0.476	0.252	0.463	0.259
UNL-85	0.429	0.145	0.489	0.295	0.487	0.312
CYC-65	0.430	0.139	0.455	0.348	0.538	0.427
CYC-85	0.432	0.144	0.493	0.388	0.548	0.483
IN-CYC-65	0.356	0.265	0.481	0.392	0.539	0.390
IN-CYC-85	0.377	0.222	0.490	0.382	0.518	0.441
PINCH-65	0.396	0.146	0.478	0.327	0.528	0.396
PINCH-85	0.429	0.141	0.504	0.354	0.552	0.436
MIX-65	0.393	0.172	0.411	0.332	0.420	0.355
MIX-85	0.408	0.159	0.420	0.330	0.498	0.401

**Table 7**  
Logarithmic standard deviations ( $\sigma$ ) of the fragility curves fitted using the different constitutive models for the frame CM2.

	IDR = 1%		IDR = 3%		IDR = 5%	
	PGA [g]	Sa( $T_1,5\%$ ) [g]	PGA [g]	Sa( $T_1,5\%$ ) [g]	PGA [g]	Sa( $T_1,5\%$ ) [g]
REF	0.393	0.262	0.526	0.368	0.597	0.470
UNL-65	0.390	0.257	0.532	0.383	0.576	0.450
UNL-85	0.390	0.257	0.532	0.383	0.576	0.450
CYC-65	0.393	0.251	0.445	0.344	0.547	0.402
CYC-85	0.395	0.252	0.491	0.368	0.592	0.461
IN-CYC-65	0.370	0.348	0.450	0.450	0.447	0.502
IN-CYC-85	0.379	0.315	0.394	0.376	0.425	0.421
PINCH-65	0.397	0.246	0.456	0.351	0.489	0.305
PINCH-85	0.387	0.257	0.515	0.360	0.595	0.469
MIX-65	0.370	0.346	0.509	0.476	0.540	0.498
MIX-85	0.369	0.315	0.479	0.443	0.526	0.446

following capacity design rules shows a median capacity in terms of Sa ( $T_1$ ) about 15% larger than CM1, for which the strong column-weak beam criterion was not used in design. On the other hand, the standard deviation values are similar.

Finally, as expected, it is worth noticing that the variability for fragilities in terms of PGA is higher than the dispersion of curves in terms of Sa( $T_1$ ), because of the higher efficiency of this latter ground motion intensity measure [59]. It is also interesting to highlight that increasing the IDR threshold, and therefore the extent of inelastic deformations, the total variance tends to increase. It should be noticed that, in the definition of the global fragility curves, every model has the same likelihood as being the true-model. As discussed in Section 6 it is

possible to consider importance weights, which could be defined based on expert opinion, for the various models by associating weight coefficients to the different terms of the likelihood function in Eq. (4).

## 6. Analysis of variance

To evaluate the contributions to the total variance of the structural capacity values, related to record-to-record variability and model uncertainty, Analysis Of Variance (ANOVA) was used [61]. The ANOVA procedure used herein is based on a 2-way crossed classification model. According to this model the general observation of structural capacity  $y_{i,p}$  either in terms of PGA or Sa( $T_1$ ), can be written as:

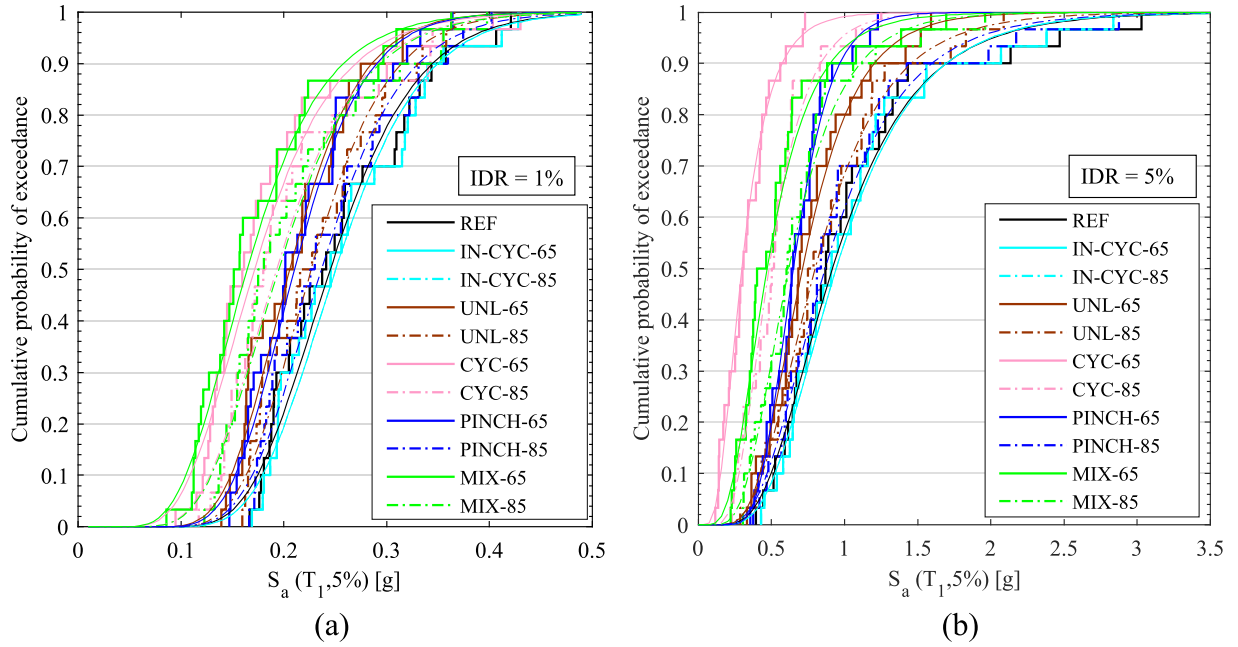


Fig. 8. Empirical cumulative fragility curves and best fitting lognormal models for the different constitutive models for the frame CM2: (a) IDR = 1% and (b) IDR = 5%.

$$\ln(y_{i,j}) = \mu + \gamma_i + \chi_j + \eta_{i,j} \quad (5)$$

where  $\mu$  is a general mean;  $\gamma_i$  is the effect of the  $i$ -th ground-motion (i.e.  $y_{i,j}$  is obtained using the ground-motion record  $i$ ),  $\chi_j$  is the effect of the  $j$ -th hysteretic model and  $\eta_{i,j}$  represents their interaction. The model in Eq. (5) is a random effects model which assumes that  $\chi$  - ,  $\gamma$  - and  $\eta$  - effects are random with zero means, variances  $\sigma_\chi^2$ ,  $\sigma_\gamma^2$  and  $\sigma_\eta^2$ , respectively, and all covariance terms equal to zero. These assumption represent the customary formulation of random-effects in random or mixed models [61]. Assuming that the aforementioned factors are normal it is possible to write the distribution of the logarithmic structural capacity as:

$$\ln(Y) \sim \text{Normal}(\mu, \sqrt{\sigma_\chi^2 + \sigma_\gamma^2 + \sigma_\eta^2}) \quad (6)$$

It should be noticed that Eq. (6) represents a simple analytical fragility model, similar to Eq. (3) [60,62]. The three components of the variance, for a balanced design of the numerical simulations (i.e. each ground-motion is used with each constitutive model), can be simply estimated as described in [56]. In particular, the analysis of variance table for the 2-way crossed classification model under consideration can be defined as shown in Table 9, where the different means are defined as follows:

$$\bar{y}_{..} = \frac{1}{g \cdot c} \sum_{i=1}^g \sum_{j=1}^c y_{ij}, \quad \bar{y}_{i.} = \frac{1}{c} \sum_{j=1}^c y_{ij}, \quad \bar{y}_{.j} = \frac{1}{g} \sum_{i=1}^g y_{ij} \quad (7)$$

where  $g$  indicates the number of ground-motion records used, and  $c$  the number of constitutive relationships. The ANOVA estimators of variance components can then be computed as follows:

$$\begin{aligned} \hat{\sigma}_\gamma^2 &= \frac{MSA - MSAB}{b} \\ \hat{\sigma}_\chi^2 &= \frac{MSB - MSAB}{a} \\ \sigma_\eta^2 &= MSAB \end{aligned} \quad (8)$$

The model parameters can also be estimated using the maximum likelihood method. The likelihood function, under normality assumptions, can be written as:

$$L(\mu, \mathbf{V} | \ln(\mathbf{y})) = \frac{\exp\left(-\frac{1}{2}(\ln(\mathbf{y}) - \mu\mathbf{1})^T \mathbf{V}^{-1}(\ln(\mathbf{y}) - \mu\mathbf{1})\right)}{(2\pi)^{\frac{1}{2}N} |\mathbf{V}|^{\frac{1}{2}}} \quad (9)$$

where  $\mathbf{y}$  indicates a vector containing the observed structural capacity values and  $\mathbf{V}$  is the variance covariance matrix, defined as:

$$\mathbf{V} = (\mathbf{I}_\chi \otimes \mathbf{1}_\gamma) \sigma_\chi^2 + (\mathbf{1}_\chi \otimes \mathbf{I}_\gamma) \sigma_\gamma^2 + (\mathbf{I}_\chi \otimes \mathbf{I}_\gamma) \sigma_\eta^2 \quad (10)$$

The symbol  $\otimes$  indicates the Kronecker product,  $\mathbf{I}_\chi$  and  $\mathbf{I}_\gamma$  are 11x11 and 30x30 identity matrices, respectively, and  $\mathbf{1}_\chi$ ,  $\mathbf{1}_\gamma$  11x1 and 30x1 unit vectors. The estimates of the parameters of the model, i.e.  $\mu$ ,  $\sigma_\chi^2$ ,  $\sigma_\gamma^2$ ,  $\sigma_\eta^2$  are the values that maximize the likelihood function (or its logarithm, which is easier to manage from a computational point of view). Because of the nature of the variance-covariance matrix which is sparse, specific computational methods are required for the solution of the maximization problem [61]. The model described so far can be further extended in order to consider importance weights for the various constitutive model. These weights could be defined based on expert opinion and engineering judgment. In order to include them in the fitting procedure, the easiest approach that can be followed is resampling structural capacity data proportionally to the importance weights, thus increasing the number of terms in the likelihood function. As an example, the weights reported in Table 10, based on engineering judgment, were used. It should be noticed that the judgement of different experts could be easily combined using logic tree.

Tables 11 and 12 report the results obtained for models CM1 and CM2, respectively. As expected, for low IDR values (1%) the model uncertainty has a marginal role because inelastic deformations are limited. On the contrary, for IDR values of 3% and 5%, the uncertainty related to constitutive model definition plays an important role, with variance contributions comparable to the record-to-record variability. For example considering IDR = 5% and  $S_a(T_1)$  these contributions represent 38% and 42% of the total variance for models CM1 and CM2, respectively. The contribution due to the interaction between model uncertainty and record-to-record variability is 8% for CM1 and 17% for CM2. This contribution represents the dependency of record to record variability on the constitutive model, as discussed in Section 5. For these levels of IDR, at which the structures show significant inelastic deformations, the choice of the constitutive model plays a crucial role in influencing the dynamic capacity of buildings. Therefore, when a

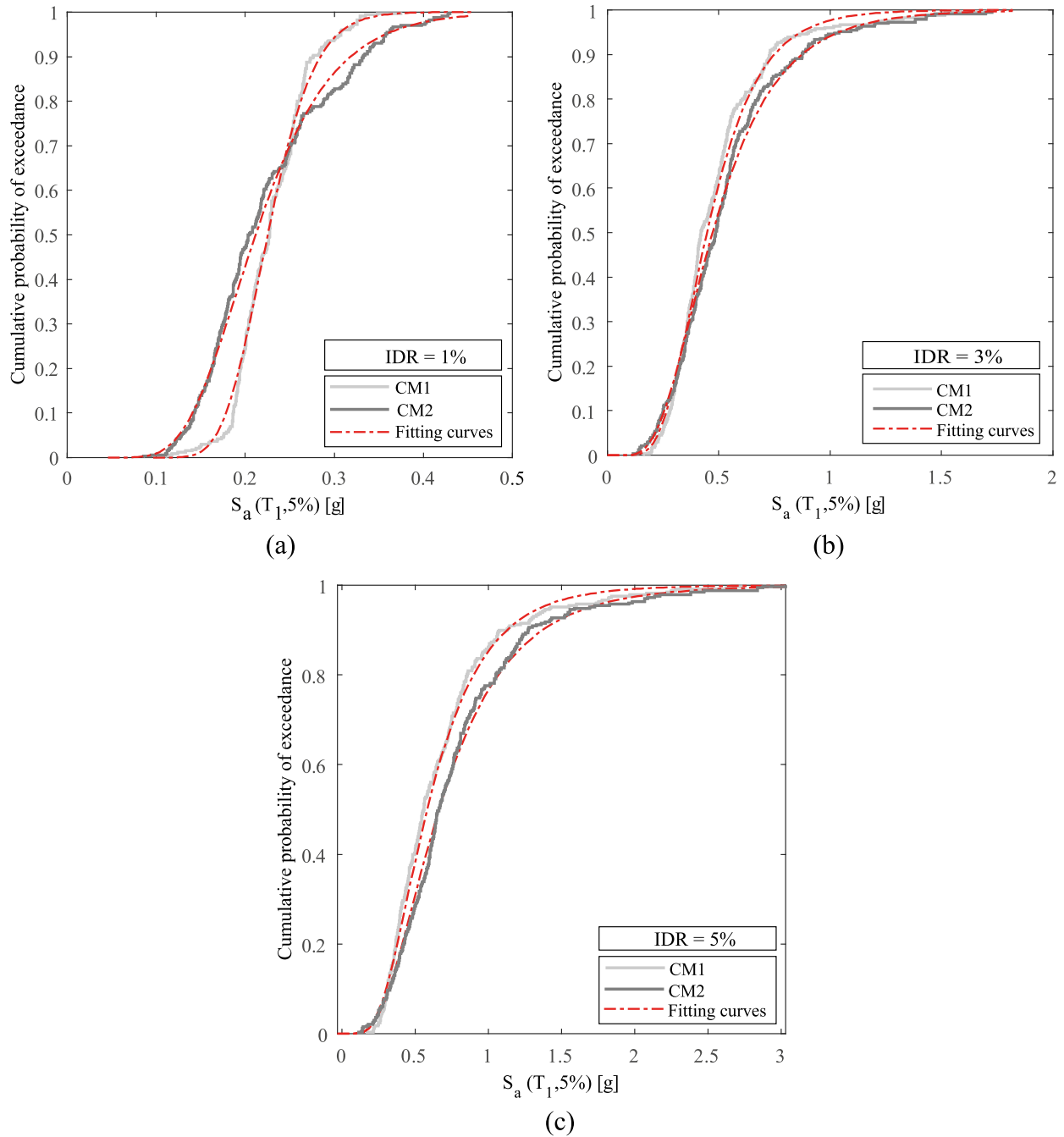


Fig. 9. Global empirical fragility curves and best fitting lognormal curves obtained for (a) IDR = 1%, (b) IDR = 3% and (c) IDR = 5%.

Table 8

Median values and logarithmic standard deviations of the global lognormal fragility models the three levels of IDR considered in the study, for the structures CM1 and CM2, in terms of both PGA and spectral acceleration  $S_a(T_1, 5\%)$ .

Parameter	IDR = 1%		IDR = 3%		IDR = 5%	
	PGA [g]	$S_a(T_1, 5\%)$ [g]	PGA [g]	$S_a(T_1, 5\%)$ [g]	PGA [g]	$S_a(T_1, 5\%)$ [g]
CM1						
exp( $\mu$ )	0.2363	0.2252	0.4685	0.4469	0.6124	0.5829
$\sigma$	0.4161	0.1821	0.5210	0.4036	0.6109	0.5150
CM2						
exp( $\mu$ )	0.2218	0.2115	0.4935	0.4719	0.6911	0.6633
$\sigma$	0.4105	0.3185	0.5567	0.4798	0.6402	0.5665

**Table 9**  
Analysis of variance table for a two way crossed classification model with balanced data.

Source of variation	Degrees of freedom	Sum of squares	Mean Square
$\gamma$ - factor	$g - 1$	$SSA = c \sum_{i=1}^g (\bar{y}_{i.} - \bar{y}_{..})^2$	$MSA = \frac{SSA}{g - 1}$
$\chi$ - factor	$c - 1$	$SSB = g \sum_{j=1}^c (\bar{y}_{.j} - \bar{y}_{..})^2$	$MSB = \frac{SSB}{c - 1}$
$\eta$ - factor	$(g - 1)(c - 1)$	$SSAB = \sum_{i=1}^g \sum_{j=1}^c (y_{ij} - \bar{y}_{i.} - \bar{y}_{.j} + \bar{y}_{..})^2$	$MSAB = \frac{SSAB}{(g - 1)(c - 1)}$
Total	$g \cdot c - 1$	$SST = \sum_{i=1}^g \sum_{j=1}^c (y_{ij} - \bar{y}_{..})^2$	

**Table 10**  
Importance weights considered for the different constitutive models.

	CM1	CM2
REF	0.4	0.4
UNL-65	0.5	0.4
UNL-85	0.4	0.8
CYC-65	0.8	0.5
CYC-85	0.6	0.8
IN-CYC-65	0.8	0.4
IN-CYC-85	0.6	0.8
PINCH-65	0.8	0.5
PINCH-85	0.6	0.5
MIX-65	1.0	0.7
MIX-85	0.6	1.0

**Table 11**  
Variance and standard deviation components for the frame CM1. Numbers in brackets indicate the percentage of the total variance.

	IDR = 1%		IDR = 3%		IDR = 5%	
	PGA	PSA (T <sub>1</sub> )	PGA	PSA (T <sub>1</sub> )	PGA	PSA (T <sub>1</sub> )
$\sigma_\gamma^2$	0.161 (84%)	0.023 (70%)	0.209 (76%)	0.102 (63%)	0.245 (65%)	0.131 (50%)
$\sigma_\chi^2$	0.024 (13%)	0.004 (12%)	0.050 (18%)	0.047 (29%)	0.113 (30%)	0.109 (42%)
$\sigma_\eta^2$	0.006 (3%)	0.006 (18%)	0.015 (5%)	0.014 (9%)	0.021 (6%)	0.022 (8%)
$\sigma_{TOT}^2$	0.191	0.033	0.274	0.163	0.379	0.262
$\sigma_\gamma$	0.401	0.153	0.457	0.319	0.495	0.362
$\sigma_\chi$	0.067	0.066	0.223	0.218	0.336	0.331
$\sigma_\eta$	0.080	0.079	0.124	0.120	0.145	0.149
$\sigma_{TOT}$	0.437	0.182	0.523	0.404	0.616	0.512

**Table 12**  
Variance and standard deviation components for the frame CM2. Numbers in brackets indicate the percentage of the total variance.

	IDR = 1%		IDR = 3%		IDR = 5%	
	PGA	PSA (T <sub>1</sub> )	PGA	PSA (T <sub>1</sub> )	PGA	PSA (T <sub>1</sub> )
$\sigma_\gamma^2$	0.141 (83%)	0.075 (72%)	0.204 (65%)	0.127 (54%)	0.235 (56%)	0.148 (45%)
$\sigma_\chi^2$	0.021 (12%)	0.021 (20%)	0.076 (24%)	0.077 (33%)	0.121 (29%)	0.124 (38%)
$\sigma_\eta^2$	0.008 (5%)	0.008 (8%)	0.033 (11%)	0.030 (13%)	0.062 (15%)	0.057 (17%)
$\sigma_{TOT}^2$	0.17	0.104	0.313	0.234	0.418	0.329
$\sigma_\gamma$	0.376	0.273	0.452	0.356	0.485	0.384
$\sigma_\chi$	0.146	0.145	0.276	0.278	0.348	0.353
$\sigma_\eta$	0.087	0.087	0.180	0.173	0.248	0.238
$\sigma_{TOT}$	0.412	0.322	0.559	0.484	0.647	0.574

unique hysteretic constitutive model cannot be identified for a given structure, the introduction of the corresponding epistemic uncertainty, significantly increases the total variance of structural capacity.

**7. Final remarks**

The recent understanding of the importance of considering model uncertainties, especially in the definition of seismic collapse capacity has led the scientific community to focus their attention on this source of variability. In this context, the present paper mainly concerns the evaluation of the contribution to the total variability of fragility models, related to epistemic uncertainty on the choice of a proper hysteretic model. Two alternative designs, with and without capacity design rules, for a case study frame building, and eleven alternative constitutive models, based on the typical experimental behaviour of RC elements were considered in the study.

By means of IDAs we evaluated fragility for different IDR thresholds (i.e. 1%, 3% and 5%), considering either PGA or Sa(T<sub>1</sub>) as ground-motion intensity measures. These data were analysed using ANOVA in order to identify and split the variance contribution due to record-to-record variability and model-to-model variability. We found that the variability contribution for model-to-model uncertainty is relevant for IDR values larger than 1% and it increases as the extent of inelastic deformation enlarges. At the collapse condition (IDR = 5%), magnitude of model-to-model variability and record-to-record variability are similar. Therefore, when the hysteretic constitutive model cannot be identified for a given structure, the introduction of the corresponding epistemic uncertainty, significantly increases the total variance of structural capacity.

**Appendix A. Supplementary material**

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.engstruct.2019.03.064>.

**References**

- [1] Cornell CA, Krawinkler H. Progress and challenges in seismic performance assessment. PEER Center News, Spring; 2000. <http://peer.berkeley.edu/news/2000spring/index.html>.
- [2] Der Kiureghian A. Non-ergodicity and PEER's framework formula. *Earthquake Engng Struct Dyn* 2005;34:1643–52.
- [3] Liel AB, Haselton CB, Deierlein GG, Baker JW. Incorporating modeling uncertainties in the assessment of seismic collapse risk of buildings. *Struct Saf* 2009;31(2):197–211.
- [4] Cornell CA, Vanmarcke EH. The major influences on seismic risk. In: 3rd World conference on earthquake engineering, Santiago, Chile, A-1; 1969. p. 69–83.
- [5] McGuire RK, Shedlock KM. Statistical uncertainties in seismic hazard evaluations in the United States. *Bull. Seism Soc. Am.* 1981;71:1287–308.
- [6] Kulkarni RB, Youngs RR, Coppersmith KJ. Assessment of confidence intervals for results of seismic hazard analysis. Proceedings, 8th World Conference on Earthquake Engineering, vol. 1. Tokyo, Japan: International Association for Earthquake Engineering; 1984. p. 263–70.
- [7] Abrahamson NA, Somerville PG, Cornell AC. Uncertainty in numerical strong motion predictions. In: Proc. of the 4th US National conference on earthquake engineering, Palm Springs, CA; 1991.
- [8] Budnitz RJ, Apostolakis G, Boore DM, Cluff LS, Coppersmith KJ, Cornell CA, et al. Recommendations for Probabilistic Seismic Hazard Analysis: Guidance on Uncertainty and Use of Experts. NUREG/CR-6372, 2 volumes. Washington, D.C.: U. S. Nuclear Regulatory Commission; 1997.
- [9] Abrahamson NA, Bommer JJ. Probability and uncertainty in seismic hazard analysis. *Earthquake Spectra* 2005;21:603–7.
- [10] Bommer J, Abrahamson NA. Why do modern probabilistic seismic-hazard analyses often lead to increased hazard estimates? *Bull Seism Soc Am* 2006;96:1967–77.
- [11] Bommer JJ, Scherbaum F. The use and misuse of logic-trees in probabilistic seismic hazard analysis. *Earthquake Spectra* 2008;24:997–1009.



- [12] Scherbaum F, Kuehn NM. Logic tree branch weights and probabilities: Summing up to one is not enough. *Earthquake Spectra* 2011;27:1237–51.
- [13] Musson R. On the nature of logic trees in probabilistic seismic hazard assessment. *Earthquake Spectra* 2012;28:1291–6.
- [14] Ordaz M, Arroyo D. On uncertainties in probabilistic seismic hazard analysis. *Earthquake Spectra* 2016;32(3):1405–18.
- [15] Yun SY, Hamburger RO, Cornell CA, Foutch DA. Seismic performance evaluation for steel moment frames. *J Struct Eng* 2002;128(4):12.
- [16] Vamvatsikos D, Cornell CA. Direct estimation of seismic demand and capacity of multi-degree of freedom systems through incremental dynamic analysis of single degree of freedom approximation. *J Struct Eng (ASCE)* 2005;131(4):589–99.
- [17] Han SW, Chopra AK. Approximate incremental dynamic analysis using the modal pushover analysis procedure. *Earthquake Eng Struct Dyn* 2006;35:1853–73.
- [18] Zareian F, Krawinkler H. Assessment of probability of collapse and design for collapse safety. *Earthquake Eng Struct Dyn* 2007;36:1901–14.
- [19] Kappos AJ, Chryssanthopoulos MK, Dymiotis C. Uncertainty analysis of strength and ductility of confined reinforced concrete members. *Eng Struct* 1999;21:195–208.
- [20] Kwon OS, Elashai A. The effect of material and ground motion uncertainty on the seismic vulnerability curves of RC structure. *Eng Struct* 2006;28:289–303.
- [21] Dolsek M. Incremental dynamic analysis with consideration of modeling uncertainties. *Earthquake Engng Struct Dyn* 2009;38:805–25.
- [22] Vamvatsikos D, Fragiadakis M. Incremental dynamic analysis for estimating seismic performance sensitivity and uncertainty. *Earthquake Engng Struct Dyn* 2010;39:141–63.
- [23] Ibarra L, Krawinkler H. Global collapse of frame structures under seismic excitations. Blume Center TR 152. Stanford University; 2003.
- [24] Baker JW, Cornell CA. Uncertainty propagation in probabilistic seismic loss estimation. *Struct Saf* 2008;30:236–52.
- [25] Borgonovo E, Zentner I, Pellegrini A, Tarantola S, de Rocquigny E. On the importance of uncertain factors in seismic fragility assessment. *Reliab Eng Syst Saf* 2013;109:66–76.
- [26] Celik OC, Ellingwood BR. Seismic fragilities for non-ductile reinforced concrete frames – role of aleatoric and epistemic uncertainties. *Struct Saf* 2010;32:1–12.
- [27] Ugruhan B, Baker JW, Deierlein GG. Uncertainty estimation in seismic collapse assessment of modern reinforced concrete moment frame buildings. In: 10th U.S. national conference on earthquake engineering, July 21–25, Anchorage, Alaska; 2014.
- [28] Vamvatsikos D, Cornell CA. Incremental dynamic analysis. *Earthquake Engng Struct Dyn* 2002;31:491–514. <https://doi.org/10.1002/eqe.141>.
- [29] CEN (European Committee for Standardization). “General rules and rules for buildings.” Eurocode 2, Brussels; 2004.
- [30] CEN (European Committee for Standardization). “General rules, seismic actions and rules for building.” Eurocode 8, Brussels; 2005.
- [31] McKenna F, Scott MH, Fenves GL. Nonlinear finite element analysis software architecture using object composition. *J Comput Civil Eng* 2010;24(1):95–107.
- [32] FEMA 273. NEHRP Guidelines for the Seismic Rehabilitation of Buildings, prepared by the Building Seismic Safety Council for the Federal Emergency Management Agency, Washington, D.C.; 1997.
- [33] OpenSees (Open System for Earthquake Engineering Simulation). Command manual; 2017. Website [http://opensees.berkeley.edu/wiki/index.php/Command\\_Manual](http://opensees.berkeley.edu/wiki/index.php/Command_Manual).
- [34] Nielsen NN, Imbeault FA. Validity of various hysteretic systems. In: Proceedings, of 3rd Japan national conference on earthquake engineering; 1970. p. 707–14.
- [35] Bradley BA. A framework for validation of seismic response analyses using seismometer array recordings. *Soil Dyn Earthquake Eng* 2011;31(3):512–20. <https://doi.org/10.1016/j.soildyn.2010.11.008>.
- [36] Waugh JD, Sritharan S. Lessons learned from seismic analysis of a seven-story concrete test building. *J Earthquake Eng* 2010;14(3):448–69. <https://doi.org/10.1080/13632460903206485>.
- [37] Sousa R, Correia AA, Almeida JP, Pinho R. Blind prediction tests as a benchmark to improve the seismic response of fibre models. In: Proc. of 2nd European conference on earthquake engineering and seismology, 25–29 August, Istanbul, Turkey; 2014.
- [38] Bommer JJ, Scherbaum F. The use and misuse of logic trees in probabilistic seismic hazard analysis. *Earthquake Spectra* 2008;24(4):997–1009.
- [39] PEER (Pacific Earthquake Engineering Research Center). Structural Performance Database; 2003. Website, <http://nisee.berkeley.edu/spd/index.html>.
- [40] Erberik MA. Energy based seismic assessment of degrading systems PhD Thesis Civil Engineering Department, METU; 2001
- [41] Lignos DG, Krawinkler H. Sidesway collapse of deteriorating structural systems under seismic excitation. Rep.No.TB 177. Stanford (CA): The John A. Blume Earthquake Engineering Research Center, Stanford University; 2012 [electronic version: [https://blume.stanford.edu/tech\\_reports](https://blume.stanford.edu/tech_reports)].
- [42] Elwood KJ. Shake table tests and analytical studies on the gravity load collapse of reinforced concrete frames PhD. Dissertation Berkeley: Department of Civil and Environmental Engineering, University of California; 2002.
- [43] Fardis MN, Biskinis DE. Deformation capacity of RC members, as controlled by flexure or shear. In: Proceeding of the Symposium in Honor of Professor Otani, Japan; 2003.
- [44] Sezen H. Evaluation and testing of existing reinforced concrete building columns. CE299 Report. Berkeley: University of California; 2000.
- [45] Mansour MY, Lee JY, Hindic R. Analytical prediction of the pinching mechanism of RC elements under cyclic shear using a rotation-angle softened truss model. *Eng Struct* 2005;27:1138–50.
- [46] Kurtman B. A detailed analysis for evaluation of the degradation characteristics of simple structural systems MS Thesis Ankara, Turkey: Middle East Technical University; 2007
- [47] Clough RW, Johnston SB. Effect of Stiffness degradation on earthquake ductility requirements. In: Proceedings 2nd Japan national conference on earthquake engineering; 1966. p. 227–32.
- [48] Powell GH, Allahabadi R. Seismic damage prediction by deterministic methods: concept and procedure. *Earthquake Eng Struct Dyn* 1988;16(5):719–34.
- [49] Park Y, Ang A. Mechanistic seismic damage model for reinforced concrete. *J Struct Eng ASCE* 1985;111(4):722–39.
- [50] Roufaiel MSL, Meyer C. Reliability of concrete frames damaged by earthquakes. *J Struct Eng ASCE* 1987;113(3):445–57.
- [51] Mehanny S, Deierlein G. Seismic damage and collapse assessment of composite moment frames. *J Struct Eng ASCE* 2001;127(9):1045–53.
- [52] Colombo A, Negro P. A damage index of generalised applicability. *Eng Struct* 2005;27(8):1164–74.
- [53] Sinha R, Shiradhonkar SR. Seismic damage index for classification of structural damage-closing the loop. In: 15th World conference on earthquake engineering. Lisboa, Portugal; 2012.
- [54] Buratti N, Stafford PJ, Bommer JJ. Earthquake accelerogram selection and scaling procedures for estimating the distribution of drift response. *J Struct Eng* 2011;137(3):345–57. [https://doi.org/10.1061/\(asce\)st.1943-541x.0000217](https://doi.org/10.1061/(asce)st.1943-541x.0000217).
- [55] Baker JW, Cornell CA. Spectral shape, epsilon and record selection. *Earthquake Eng Struct Dyn* 2006;35(9):1077–95.
- [56] Baker JW. Conditional mean spectrum: tool for ground motion selection. *J Struct Eng* 2011;137(3):322–31.
- [57] Bradley BA. A generalized conditional intensity measure approach and holistic ground-motion selection. *Earthquake Engng Struct Dyn* 2010;39:1321–42.
- [58] Fajfar P, Dolsek M, Marusic D, Stratan A. Pre- and post-test mathematical modelling of a plan-asymmetric reinforced concrete frame building. *Earthquake Eng Struct Dyn* 2006;35:1359–79.
- [59] Tothong P, Luco N. Probabilistic seismic demand analysis using advanced ground motion intensity measures. *Earthquake Eng Struct Dyn* 2007;36(13):1837–60.
- [60] Buratti N, Minghini F, Ongareto E, Savoia M, Tullini N. Empirical seismic fragility for the precast RC industrial buildings damaged by the 2012 Emilia (Italy) earthquakes. *Earthquake Engng Struct Dyn* 2017;46:2317–35. <https://doi.org/10.1002/eqe.2906>.
- [61] Searle SR, Casella G, McCulloch CE. Variance components. Wiley Ed; 2006. ISBN: 978-0-470-00959-8.
- [62] Buratti N, Ferracuti B, Savoia M. Response Surface with random factors for seismic fragility of reinforced concrete frames. *Struct Saf* 2010;32(1):42–51.