ORIGINAL ARTICLE

# Parental breeding decisions and genetic quality predict social structure of independent offspring

Victoria R. Franks<sup>1</sup> | Rose Thorogood<sup>2</sup> | Patricia Brekke<sup>3</sup>

<sup>1</sup>Department of Biological Sciences, University of Chester, Chester, UK <sup>2</sup>Helsinki Institute of Life Science (HiLIFE), University of Helsinki, Helsinki, Finland <sup>3</sup>Institute of Zoology, Zoological Society of London, London, UK

#### Correspondence

Victoria R. Franks, Department of Biological Sciences, University of Chester, Parkgate Road, Chester CH1 4BJ, UK. Email: v.franks@chester.ac.uk

Funding information Association for the Study of Animal Behaviour; Research England

Handling Editor: Andrew DeWoody

### Abstract

Across the animal kingdom, newly independent juveniles form social associations that influence later fitness, mate choice and gene flow, but little is known about the ontogeny of social environments, particularly in wild populations. Here we test whether associations among young animals form randomly or are influenced by environmental or genetic conditions established by parents. Parents' decisions determine natal birth sites, which could affect who independent young initially encounter; secondly, mate choice determines genetic condition (e.g. inbreeding) of young and the parental care they receive, which can affect sociability. However, genetic and environmental factors are confounded unless related offspring experience different natal environments. Therefore, we used a long-term genetic pedigree, breeding records and social network data from three cohorts of a songbird with high extra-pair paternity (hihi, Notiomystis cincta) to disentangle (1) how nest location and relatedness contribute to association structure once juveniles disperse away from birth sites, and (2) if juvenile and/or parental inbreeding predicts individual sociability. We detected positive spatial autocorrelation: hihi that fledged closer by were more likely to associate even after dispersing, irrespective of genetic relatedness. Juvenile inbreeding did not predict sociability, but those raised by more inbred fathers formed more, stronger, associations, which did not depend on whether that male was the genetic parent or not. These results suggest that the natal environment created by parents, rather than focal genetic condition, establishes the foundation for social associations. Overall, we highlight how social inheritance may play an important role in population dynamics and evolutionary potential in wild animals.

#### KEYWORDS

inbreeding, juvenile, natal environment, parents, relatedness, social structure

### 1 | INTRODUCTION

Early independence represents a uniquely important, and yet understudied, point in the development of social environments in wild animals. As young leave their parents, they encounter their first opportunities to interact with other individuals outside their immediate family group. Social interactions during this sensitive early-life period have a suite of consequences for later life behaviour and fitness (Cantor et al., 2021), including sociality (Brandl et al., 2019), senescence (Péron et al., 2010; Turner et al., 2021) and reproductive

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strategies (Macario et al., 2017; Turner et al., 2021). Furthermore, when the structure of the social environment in newly-independent juveniles determines these interaction opportunities they have the potential to delineate boundaries for crucial processes including gene flow (Sugg et al., 1996) and cultural and social evolution (Cantor et al., 2021; Kuijper & Johnstone, 2019). However, despite a growing body of evidence showing that individuals do not interact equally (Sosa et al., 2021; Wey et al., 2008), we have little knowledge of the factors that contribute to social structure in free-living young animals. In part this is because juveniles are often more free-ranging, cryptic and difficult to study in comparison to when they are either dependent young or adults with breeding territories or home ranges.

The genetic relatedness of associating young is likely to be a fundamental factor underpinning the fitness consequences of sociality (Wittemyer et al., 2009). Developing associations with genetically dissimilar individuals can reduce competition between relatives and the likelihood of inbreeding (Godfrey et al., 2014; Hirsch et al., 2013; Kurvers et al., 2013; Mourier & Planes, 2021), particularly if social structure predicts later mating decisions (Firth & Sheldon, 2016). On the other hand, associating with kin (e.g. Archie et al., 2006; Carter et al., 2013; Kurvers et al., 2013) can enhance social learning or cooperation (Kerth, 2008; Schwab et al., 2008) and improve growth, survival and indirect fitness (Chakrabarti et al., 2020; Feh, 1999; Gerlach et al., 2007; Thünken et al., 2016). Finally, social learning can also be enhanced by associating with familiar individuals, regardless of their genetic relatedness, as they are expected to be more reliable demonstrators (Camacho-Alpízar & Guillette, 2023). Disentangling whether familiarity or kin are favoured in social interactions in wild populations is however challenging as natal environment is usually conflated with genetic relatedness. Additionally, other factors such as philopatry may also lead siblings to continue to associate rather than it being a result of selecting for genetic relatedness itself (Leedale et al., 2020). Thus, whether juveniles associate with kin as a direct response to relatedness or associate simply due to their shared natal environment is currently unclear.

Alongside influencing who associates with whom, genetics can also impact on sociality at the individual level via traits that have a strong genetic basis: for example, larger individuals may succeed in gaining more associates than smaller conspecifics (Pack et al., 2009). However, genetic effects may also be more indirect. The genotypes of parents can affect traits in their offspring via the environment they create during rearing (Kong et al., 2018), known as 'genetic nurturing'. For example, if parents create a poorer or more stressful early-life environment this can in turn affect their offspring's later life social strategies (Boogert et al., 2014; Brandl et al., 2019; Farine, Spencer, & Boogert, 2015). One key aspect of genetics that affects both individual traits and parental care, and may therefore impact on sociality either directly or indirectly, is inbreeding. First, being more inbred alters an individual's life history (e.g. condition and secondary sexual characteristics; Bolund et al., 2010) and behaviour [e.g. personality (Herdegen-Radwan, 2019; Müller & Juškauskas, 2018), dispersal strategies (Daniels & Walters, 2000) and cooperation (Wells et al., 2020)]. These traits can affect both quality and quantity of associations: as one

example, animals with more reactive personalities have fewer and less stable social network connections (Aplin et al., 2013). Second, inbreeding alters parents' investment strategies in their offspring (Duthie et al., 2016; Pooley et al., 2014; Wells et al., 2020), and can therefore alter early-life rearing conditions (Pooley et al., 2014). While inbreeding depression is not always apparent in benign environments (Armbruster & Reed, 2005; Crnokrak & Roff, 1999), early independence is one of the most challenging life stages and therefore the consequences for inbred individuals may be more likely to be detected (Duntsch et al., 2023). However, while this demonstrates how traits and states resulting from inbreeding may be linked to juvenile social behaviour, there have been no studies explicitly testing the relationship between sociality and different generational levels of inbreeding per se.

While the costs and benefits of social living can have a strong genetic basis, this component cannot be considered in isolation. The need to consider environmental context when studying animal sociality has become increasingly recognized in recent years (Albery, Kirkpatrick, et al., 2021; Albery, Morris, et al., 2021; Evans et al., 2020; Sosa et al., 2021; Spiegel & Pinter-Wollman, 2020; Webber & Vander Wal, 2018). Abiotic factors such as habitat configuration and resource distribution shape how animals interact with a landscape and in turn affect association opportunities (He et al., 2019). For juveniles, natal site may be a vital environmental factor in shaping their first non-family associations as they disperse away from their natal territory, because juveniles born in territories in close proximity could experience similar habitat effects on their movement and settlement decisions (Bowler & Benton, 2005; Fronhofer et al., 2018). Conversely, if kin disperse away from natal sites to avoid inbreeding (Bowler & Benton, 2005), then we might expect weaker associations between juveniles from similar locations. Ultimately, if natal site acts as a physical factor impacting on social structure, then juvenile association opportunities may be determined even before they are born through breeding and settlement decisions made by their parents (Ilany & Akçay, 2016).

The likelihood that juveniles will co-occur in space also varies temporally (Pasquaretta et al., 2021; Psorakis et al., 2015; Whitehead & Dufault, 1999). During the breeding season, variation in resources at global and local scales can create variation in the timing of mating, gestation and eventually when young become independent (Ringsby et al., 2002). Thus, if juveniles become independent at the same time, they may be more likely to encounter each other, particularly in combination with spatial effects or if there are also age-associated preferences in habitat use (Ferrer & Penteriani, 2003) or social tendencies (Turner et al., 2017). However, explicitly testing for combined spatiotemporal effects is often largely overlooked in studies of animal sociality (He et al., 2019; Sosa et al., 2021; Wey et al., 2015). The effects of relatedness, environment and timing are often confounded because siblings from the same clutch or litter are born in the same place at the same time, and so understanding the individual contributions of each component to natural social structure is often impossible. Furthermore, teasing apart these effects also requires a high level of detailed breeding and genetic data, which are rarely available from wild animal populations (Clutton-Brock & Sheldon, 2010). As

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a result, multivariate analysis of complex combined factors contributing to animal social structure (such as space, time and relatedness) is challenging, and few studies have begun to tackle this area in detail (Sosa et al., 2021; Wolf & Trillmich, 2008).

While disentangling the contributions of genetics, space and time to early-life social structure is challenging in wild populations, some species do provide a natural opportunity to unpick these relationships. The hihi (Notiomystis cincta), a threatened Aotearoa/ New Zealand passerine, has one of the highest known rates of avian extra-pair paternity (mean frequency of EPP in broods reported as  $0.68 \pm 0.012$ ; Brekke et al., 2013), meaning that not all nestmates are closely related and close kin are not only within nests. In this study, we combined this natural phenomenon with a wealth of detailed genetic and breeding data available on our focal study population (Tiritiri Matangi Island, Aotearoa) to provide a rare insight into the factors contributing to early-life social environment structure in wild animals. On Tiritiri Matangi, hihi nest in boxes that we provide has enabled all breeding attempts to be monitored and recorded since the population was established in 1995. Blood sampling of all individuals has occurred since 2005 for genetic analysis, creating a long-term pedigree with derived inbreeding and relatedness metrics (Brekke et al., 2015). Inbreeding levels in this population are relatively high ( $F_{ROH} \sim 0.29$ ; Duntsch et al., 2023), but are similar to other threatened populations that have been through bottlenecks as a result of population declines; the extent of this inbreeding has been shown to affect juvenile survival in hihi (Duntsch et al., 2023). Finally, once juvenile hihi (i.e. offspring from the current breeding season) disperse from their natal nests, they congregate at 2-3 spatially separate sites for approximately 4 months each year (Franks, Ewen, McCready, Rowcliffe, et al., 2020). While social interactions in young hihi at this stage have important influences on behaviour (Franks, Ewen, McCready, & Thorogood, 2020) and survival (Franks, Andrews, Ewen, McCready, et al., 2020), underlying social structure had not been examined at a detailed genetic or spatiotemporal level prior to this study.

Here, we used social network analysis to analyse group- and individual-level social structure, which has become a widely applicable tool to quantify non-uniform associations between animals (Krause et al., 2009; Sosa et al., 2021; Wey et al., 2008). First, we investigated if associations between juvenile hihi in each group site were stronger if birds were more related and had fledged from nest boxes that were closer together in time and/or space. Second, we tested the contribution of genetic factors to individual-level sociality by analysing whether juveniles' number of associates depended on their own or their parent's inbreeding.

### 2 | MATERIALS AND METHODS

### 2.1 | Study population

Our study was conducted over 3 years (2015–2017) on Tiritiri Matangi Island (Aotearoa/New Zealand, 36°36′00″S 174°53′21″E). The study site is a 2.5-km<sup>2</sup> island covered in a mixture of native

subtropical rainforest and more open grassland. Hihi on Tiritiri Matangi are monitored intensively during the breeding season, which is typically from early September to February. Each year, the female and the social male hold a territory and raise altricial young (up to three clutches) in the nest boxes we provide. Both parents contribute to feeding offspring once hatched, similar to other passerines: while young are nestlings, the female is the primary care giver but is assisted by the social male in provisioning; once chicks fledge, the male then becomes largely solely responsible for care before the young disperse (Castro et al., 1996). Nest boxes are checked following an established protocol which allows us to record the outcomes of all nesting attempts and identify the dam and social male caring for each brood. Blood samples are taken from all nestlings at 21 days old (hatch day = day 0) to establish genetic sires and maintain a genetically resolved pedigree. Nestlings are also ringed with unique combination of coloured leg rings at this age. Young hihi fledge at around 28 days old, and after approximately 2 weeks of post-fledging dependency they disperse from their natal territory to congregate at 2-3 spatially separate sites in forested gullies for approximately 4 months (Franks, Ewen, McCready, Rowcliffe, et al., 2020), at which time they moult into their adult plumage.

During the 3 years of our study, the hihi population varied between 180 and 270 individuals, with similar proportions of juveniles and adults (second year or older) each year. The first year of our study (2015) was a poorer breeding season than 2016 and 2017 (2015: 89 fledglings; 2016: 132 fledglings; 2017: 151 fledglings), which we accounted for in analyses. There were two juvenile group sites in the 2015 season, and three group sites in both 2016 and 2017 (Figure 1).

### 2.2 | Ethical note

Ethical approval for the observations of juvenile groups was issued through the Zoological Society of London Ethics Committee (UK). Sampling and research permits, including yearly monitoring of breeding populations following an established protocol (including nest box monitoring, blood sampling and colour-ring application) were obtained from the New Zealand Department of Conservation (permit numbers WE/32213/FAU, 557 WE/246/RES, 36186-FAU, 15073-RES, 24128-FAU, 13939-RES and 44300-FAU).

### 2.3 | Social network data

Each year from mid-January to April, we observed juvenile hihi as they congregated at group sites (15 observation sessions per site in 2015, 25 sessions per site in 2016 and 2017). Each observation session lasted 1h, which was sub-divided into 30-s time blocks (120 blocks). In every 30-second block, we recorded the coloured leg-ring combinations of all hihi perched within a 15-m radius of the observer. Any bird present across multiple blocks was re-recorded at the start of each block, so we could determine continued presence to the nearest

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FIGURE 1 Map showing the location of juvenile hihi grouping sites (white stars: 2015=red; 2016=blue; 2017=yellow) and nest boxes (white circles) included in this study. Map data: Google, Maxar Technologies, TerraMetrics, CNES/ Airbus (2022).

30s. This system was used as the most fine-scale timing possible to capture changes in the presence of small forest passerines while also being long enough to allow for identification and recording of individuals. All observations were made with binoculars (Zeiss Conquest® HD  $8 \times 42$ ) by one observer (VRF). Observations were made from the same point in each group site for each observation session, and sessions were distributed evenly across the 3months each year.

We constructed an undirected weighted social network for each cohort using the R package *asnipe* (Farine, 2013), which defined associations based on spatiotemporal proximity in our time-stamped observations of juveniles; thus, we assumed individuals seen often in the same location at similar times were more familiar (gambit of the group; Whitehead, 2008). Each association value represented how frequently each pair of juveniles (dyad) associated (from 0, never seen together to 1, always seen together). Details on each cohort, including total numbers of individuals included in each network and number of network associations, are summarized in Table S1.

### 2.4 | Breeding data

### 2.4.1 | Relatedness and inbreeding coefficients

A pedigree has been maintained in this population since 1995, with paternal resolution available from 2005 when molecular markers became available (Brekke et al., 2015). Maternity is determined during nest monitoring by identifying the female that broods each clutch and cares for the chicks. Paternity is determined using genetic information from the blood samples obtained from hihi nestlings. Genomic DNA is extracted and screened using 18 neutral microsatellite loci (15 species specific and three isolated from other passerines) that are widely distributed across the genome (complete methods previously published; Brekke et al., 2013).

Using the pedigree, we calculated relatedness (r) between all dyads in each social network using the R package *nadiv* (Wolak, 2012). Relatedness ranged from 0 (no common ancestors) to approximately 0.5 (full siblings), though exact relatedness values could be slightly higher due to historical inbreeding (maximum r=0.7). Pedigree-derived inbreeding coefficients (F, the expected proportion of an individual's genome that is identical by descent) were also calculated using the *nadiv* package. Relatedness among individuals and inbreeding values were both estimated using a minimum of six known ancestors (parents and all grandparents) to reduce metric bias due to pedigree depth (maximum pedigree depth for juveniles in this study=13 generations; Villemereuil et al., 2019). Any birds which did not have known parents and grandparents were removed from further analyses.

### 2.4.2 | Natal nest box location and fledging synchrony

The latitude and longitude for all nest boxes (distributed across the 2.5 km<sup>2</sup> island) were recorded as part of the ongoing monitoring of this population (Figure 1). We defined the distance (in km) between the nest boxes of every juvenile dyad in each social network using the *geodist* 

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R package (Padgham & Sumner, 2021), which calculated geodesic distance based on shortest path length between coordinates. The resulting matrix of distances provided a spatial layout of proximities between the natal nest boxes of each pair of juveniles (from 0 to 1.55 km).

Finally, a fledging date was recorded for each clutch as part of monitoring of all breeding attempts on Tiritiri Matangi: this was the date the last chick left the nest (i.e. active nest boxes near to fledging were checked daily until they were found empty). Therefore, for our analyses each juvenile within a clutch was assigned the same fledge date. We then calculated the number of days between fledge dates for every pair of juveniles in each cohort's network to give a 'fledge synchrony' value for every dyad.

### 2.5 | Dyadic analysis of pairwise associations

To investigate how relatedness, distance between natal nest boxes and fledge synchrony predicted associations in dyads of juvenile hihi, we fitted Bayesian logistic mixed effects model using a Markov chain Monte Carlo (MCMC) framework. We used the MCMCglmm package (Hadfield, 2010), because its multi-membership modelling capabilities allowed us to account for each individual appearing interchangeably in a dyad (i.e. a hihi could be individual A or individual B in each association in our undirected networks). For every model, we used non-informative priors and adjusted burn-in periods, iterations and thin intervals to ensure minimum autocorrelation and good convergence (determined in post hoc diagnostic checks; Hadfield, 2010). We inspected models for goodness of fit using the Deviance Information Criterion (DIC). To determine the predictors that best explained our data, we removed non-significant terms from the global model to obtain the most parsimonious model which had good fit (low DIC; Rushmore et al., 2013). In all analyses, we removed dyads that included any individuals that had been seen fewer than five times, because initial exploration for outlier association strengths in our data indicated that dyads observed fewer than this amount had the least reliable network measures.

Many juveniles never associated in our networks (Table S1) and different mechanisms could affect opportunity to associate versus strength of association when they did occur. Thus, we used a stepwise approach to investigate juveniles' (a) likelihood of association (binary response: 1=associated, 0=never associated), then (b) association strength. We first compared association likelihoods between nestmates and non-nestmates to understand whether sharing a nest was important for juvenile social structure. In a binary MCMC model, our main predictor was whether dyads had fledged from the same nest (categorical value where same nest=1, different nests = 0). Next, we considered association likelihood in juveniles from different nests who had variable nest box proximities and fledge dates and thus provided an opportunity to explore the effects of spatiotemporal variation on broad-scale network structure. Again, we used a binary response (1=associated, 0=never associated); here, the main predictors were distance between natal nest boxes, fledge synchrony and additionally relatedness between

dyads to understand whether genetic sibling recognition influenced social structure. Finally, we accounted for potential biases in association metrics (Franks et al., 2021): we specified the fewest number of observations for each pair as a predictor (i.e. if individual A was seen 10 times and individual B 20 times, minimum observations = 10), and included cohort year as a random effect as association opportunities may have varied among cohorts.

We then considered (b), variation in association strengths in dyads that had associated at least once. For all following models using association strength, we z transformed the values so that they were comparable among the three cohorts in our study. Nestmates all fledged on the same day and from the same location, so had values of 0 for nest box location and fledge synchrony, but their relatedness could vary ( $r \ge approximately 0.25$ ) due to EPP; meanwhile, non-nestmates fledged on different days from different locations, and had relatedness of approximately ≤0.25. Therefore, we investigated association patterns in same-nest and different-nest juveniles separately. To analyse how relatedness predicted association strengths between nestmates, our main predictor in our MCMC model was whether nestmate dyads were half or full siblings, which we determined from pedigree relatedness values as they showed a clear bimodal distribution (half siblings: r < 0.5; full siblings: r > 0.5; Figure S1). We also included a predictor that quantified whether juveniles originated from nests that contained only half siblings, full siblings or a mix of both sibling types: even though there were similar numbers of dyads from each nest type which provided equal opportunities for association (Table S2), this parameter allowed us to quantify if association strengths were consistent even when juveniles had the opportunity to associate with both sibling types. Finally, in a separate model we explored association strengths in juveniles from different nests, where relatedness, distance between natal nest boxes and fledge synchrony all varied. Again, the response variable was z-transformed association values, and our main predictors were relatedness, distance between natal nest boxes and fledging synchrony. Here, relatedness values could not easily be categorized into full and half siblings, so we used this variable as a continuous predictor. To account for potential biases in association values, we also included the fewest number of observations for each pair in both models analysing association strengths, and a random effect for year to examine overall patterns across the different cohorts.

### 2.6 | Inbreeding and sociality analysis

To examine whether sociality was predicted by both an individual's own level of inbreeding and the inbreeding of its social parents (i.e. the male and female providing care during chick rearing), we calculated the degree strength for each juvenile, which quantifies the number and strength of associations and is thus a measure of centrality in the network (Krause et al., 2015). We then used a generalized linear mixed effects model (GLMM) with each juvenile's degree strength as the response variable, which we z transformed so we could examine general trends across the years. Our main predictors were the juvenile inbreeding coefficients, female's coefficients and

|     |                                | Posterior |                  |       |
|-----|--------------------------------|-----------|------------------|-------|
|     | Parameter                      | mean      | 95% CI           | pMCMC |
| (a) | Intercept                      | -0.961    | –1.885 to –0.030 | .035  |
|     | Shared nest of origin          | 0.424     | 0.226 to 0.589   | .002  |
|     | Minimum number of observations | 0.005     | 0.002 to 0.007   | .002  |
| (b) | Intercept                      | -0.877    | -1.820 to 0.008  | .052  |
|     | Distance between nest boxes    | -0.145    | -0.233 to 0.048  | .004  |
|     | Minimum number of observations | 0.005     | 0.002 to 0.007   | .002  |
|     |                                |           |                  |       |

*Note*: Posterior means, 95% credible intervals and *p* values were calculated with a Bayesian logistic mixed effect approach using Markov chain Monte Carlo sampling. Non-significant parameters were removed from the final models.

male's coefficients; all inbreeding coefficients were *z* scored. These variables were checked for collinearity prior to inclusion in the model, but no issues were found. We also included an interaction between the male's inbreeding coefficient and whether they were the full genetic father or the social father only. We specified a random effect of nest box identity, because multiple fledglings originated from the same nests and shared the same parents. Finally, in the subset of juveniles which had both a social and genetic father (i.e. extra-pair off-spring), we directly compared the effects of inbreeding in these two different males on their offspring sociability (degree strength) using a separate model with genetic father's *f* value and social father's *f* values as predictors; again, the random intercept was natal nest box.

Analyses using social network metrics violate the assumptions of many statistical tests due to non-independence of data, which can lead to inflated Type-I and Type-II error rates (Farine & Whitehead, 2015). Furthermore, biases can be introduced from sampling effort or spatiotemporal variation, and need to be accounted for to ensure that results are valid. Therefore, we followed recent advances in network analysis (Farine & Carter, 2022; Weiss et al., 2021) and for all inbreeding analyses we used the double permutation procedure (Farine & Carter, 2022) to calculate the significance of our effects. Here, data stream permutations first calculate any potential deviation in network metrics due to unwanted biases (e.g. spatial or sampling effects). The corrected metrics are then used to investigate the effect of each parameter of interest, with node permutations used to calculate statistical significance. As well as using double permutations to reduce any effects of observation bias among individuals, we also removed individuals observed fewer than 20 times as these individuals had the least reliable degree strengths based on initial examination of the data.

### 3 | RESULTS

# 3.1 | Effect of relatedness, fledging times and nest box proximity on social structure

### 3.1.1 | Association likelihoods

The social network structure of juvenile hihi was predicted by characteristics of the natal environment. Associations were

0.3 Different nests Same nests Dyad nests of origin

**FIGURE 2** Likelihood of association between dyads of juveniles originating from different nests and the same nests. Values are predicted from the final model exploring association likelihood across all juveniles (Table 1a).

the most likely to form between juvenile hihi that had been nestmates: sharing a natal nest box was included in our final binary model exploring association likelihood across all juveniles (N dyads = 11,184, N birds = 171, Table 1a, Figure 2). Additionally, when we considered juveniles from different nests (N dyads = 10,924, N birds = 171), nest box proximity was included in the model that best explained association likelihood (Table 1b). Therefore, non-nestmates tended to form associations if they fledged from boxes located more closely together (Table 1b, Figure 3). By contrast, there was no evidence that association likelihood changed if juveniles had fledged at similar times, or if the dyad was more closely related, as neither fledge synchrony, relatedness nor any interaction between our three main predictors were included in the final model. Finally, there was also a minimal effect of number of observations on network structure, because minimum number of observations per dyad was included in both final models investigating association likelihood but the actual effect sizes were small (Table 1).

FIGURE 3 Likelihood of associations forming between juveniles from different nest boxes, as predicted by the model in Table 1b; (a) all predicted associations, (b) the 100 least likely and (c) the 100 most likely associations. All map data: Google, Maxar Technologies, TerraMetrics, CNES/ Airbus (2022).



### 3.1.2 | Association strengths

In contrast to our results from likelihood of association, we did not find evidence that our three main parameters (nest box proximity, fledge synchrony and relatedness) predicted association strength, as these parameters were not included in the final models for either nestmates or non-nestmates (nestmates: N dyads=2924, Nbirds=167; non-nestmates: N dyads=2842, N birds=167). In nestmates, no parameter significantly explained association strength, and the only parameter in the final model for non-nestmates was minimum number of observations, which had a negligible effect on associations (Table 2).

## 3.2 | Effect of early-life inbreeding environment on sociality

An individual's own level of inbreeding did not predict its number of network associates (degree strength; Table 3a; N=111 birds). However, a juvenile's sociability was significantly predicted by its social father's inbreeding. We found that juveniles whose fathers were more inbred (higher inbreeding coefficient) had higher degree strength (Table 3a, Figure 4a), independent of whether the father was genetically related to the juvenile or was only their social care provider (interaction between father's inbreeding coefficient and genetic status non-significant, Table 3a). This effect remained consistent even when we removed two fathers who had very high inbreeding coefficients, so these individuals were not driving the relationship between father's inbreeding and juvenile sociability TABLE 2 Predictors included in the final model analysing dyadic association strengths in juvenile hihi, for non-nestmates.

| Parameter                         | Posterior<br>mean | 95% CI          | рМСМС |
|-----------------------------------|-------------------|-----------------|-------|
| Intercept                         | 0.001             | -0.171 to 0.174 | .994  |
| Minimum number<br>of observations | 0.004             | 0.002 to 0.006  | .003  |

*Note*: Posterior mean, 95% credible interval and *p* values are calculated from a Bayesian logistic mixed effect model using Markov chain Monte Carlo sampling. Non-significant parameters were removed from the final model.

(Table S3a, Figure S2a). While there was a trend for juveniles with more inbred mothers to be less social (Table 3a, Figure 4b), this was not statistically significant in either the full dataset or with two mothers with extreme inbreeding coefficients removed (Table S3b; Figure S2b). Similarly, in extra-pair juveniles where we could directly compare the effects of both genetic and social fathers (N=70), only the social father's inbreeding predicted offspring sociability (Table 3b).

### 4 | DISCUSSION

Understanding the structure of the early-life social environment is crucial, because interactions in this period have the potential to determine behaviours through to adulthood (Slagsvold & Wiebe, 2011) and influence key population processes such as survival and reproduction (Nuñez et al., 2015; Slagsvold & Wiebe, 2011). However, we

TABLE 3 Results of linear mixed effect model analysing the effect of inbreeding on individual sociality (degree strength) in (a) all and (b) extra-pair, juvenile hihi.

|     | Parameter  | Est.   | SE    | t      | p <sub>double permutation</sub> |
|-----|--|--------|-------|--------|---------------------------------|
| (a) | Intercept  | 0.569  | 0.097 | 4.823  | -                               |
|     | Juvenile inbreeding coefficient (z score)                                      | 0.026  | 0.063 | 0.413  | .338                            |
|     | Mother's inbreeding coefficient (z score)                                      | -0.108 | 0.075 | -1.452 | .078                            |
|     | Inbreeding coefficient of nest tending male (z score)                          | 0.257  | 0.124 | 2.074  | .017                            |
|     | Within-/extra-pair father  | 0.015  | 0.139 | 0.110  | .465                            |
|     | Father's inbreeding coefficient $	imes$ within-/extra-pair father              | -0.180 | 0.143 | -1.254 | .100                            |
| (b) | Intercept  | 0.480  | 0.093 | 5.144  | -                               |
|     | Social father's inbreeding coefficient (z score)                               | 0.203  | 0.114 | 1.78   | .034                            |
|     | Genetic father's inbreeding coefficient (z score)                              | 0.185  | 0.136 | 1.357  | .104                            |
|     | Social father's inbreeding coefficient×genetic father's inbreeding coefficient | 0.168  | 0.225 | 0.746  | .230                            |

Note:  $p_{\text{double permutation}}$  is the p value resulting from the double permutation procedure. Significant effects (at p < .05) highlighted in bold.



FIGURE 4 Relationship between the degree strength of each juvenile and their parents' extent of inbreeding; (a) the effect of their social father's inbreeding score (this male was present during chick rearing, but was not always the genetic father); (b) the effect of their mother's inbreeding coefficient. Degree and inbreeding coefficients are z transformed. Line of best fit (blue line) with 95% CI calculated from model in Table 3.

often have limited opportunity to understand how different socioecological components impact on associations once juveniles leave their parents and predict the types of interactions available. In this study, we investigated the contribution of spatial, temporal and genetic factors to early-life social structure in groups of juvenile hihi. Across three cohorts of juveniles, we found that individuals from the same nest were most likely to form associations. Similarly, juveniles from different nests showed a tendency to associate when their nest boxes had been closer together. However, we found no evidence for effects of nest box proximity, synchrony in fledging timing or relatedness on finer scale association strengths. Finally, each juvenile's sociability (degree strength) was not predicted by their own inbreeding, but instead related to their parents': in particular, juveniles were more social if their social fathers (who had been providing care at the nest) were more inbred. Overall, these results highlight the dual importance of ecological and genetic components from early life in determining social interactions between juvenile animals, and demonstrate how these different factors may act at different levels within populations to create an emergent social structure from a young age.

Across the animal kingdom, the physical structure of the environment, such as resource distribution and spatial configuration,

shapes both individual and collective behavioural decisions which in turn determines when and where individuals co-exist in space and time (Mbizah et al., 2020; Pasquaretta et al., 2021; Sosa et al., 2021; Strandburg-Peshkin et al., 2017). In our study, we extend knowledge on the importance of habitat on social structure by highlighting how the physical natal environment contributes to the first opportunities that young animals have to associate once they are independent. In hihi, the location of nest boxes predicted early-life association likelihood: juveniles from the same nest were most likely to form associations. Association likelihood then declined as distance between nest boxes increases, though this effect was small which may be because the size of our study site (2.5 km<sup>2</sup>) limited the extent that associations could differentiate. Nevertheless, the combined evidence within and across nests highlights how an animal's physical environment can predict the basis of social structure when it leads some animals to associate, but not others (He et al., 2019). For juvenile animals, natal location may create association opportunities through one major process at this life stage: dispersal is a key event for determining the future distribution of juveniles in a population and can be shaped by the surrounding configuration of resources including food and suitable habitat (Kaemingk et al., 2019; Messier et al., 2012; Paradis et al., 1998). Thus, juveniles from similar natal

locations may share similar dispersal patterns, making natal location important in determining the opportunities for the ontogeny of social associations at the initial point of early-life independence.

Relatedness and familiarity are often inherently linked within siblings due to their shared raising environment and are difficult to tease apart (Leedale et al., 2020): even when studies in both wild birds and mammals have found evidence that siblings preferentially associate over unrelated individuals (Bonadonna & Sanz-Aguilar, 2012; Wolf & Trillmich, 2008), they rarely separate genetic relatedness from more simple cue-based familiarity. However, in our cohorts of wild hihi, we also had an opportunity to separate genetic effects on associations from ontogenetic familiarity because relatedness varied within and across nests due to EPP. Overall, there was a lack of evidence for genetic relatedness underlying dyadic association patterns, because more closely related individuals were not more likely to be connected in our networks. This supports a recent review that concluded genetic cues to kinship are rare in Class Aves overall, while familiarity from learned or environmental cues offer a more parsimonious explanation for associations between kin in most contexts (Leedale et al., 2020). Alongside the finding that associations were most likely between nestmates, this highlights the importance of natal environment for association opportunities, independent of any link with relatedness between associating individuals.

Social structure may be further mediated by traits and/or states that affect the number and strength of associations' individuals have (Croft et al., 2009; Farine, Montiglio, & Spiegel, 2015; Gartland et al., 2021; Pike et al., 2008). However, in our study we found no evidence that a juvenile hihi's own extent of inbreeding affected its social behaviour, despite inbreeding depressing juvenile survival (Duntsch et al., 2023). Instead, their individual sociability was predicted by their fathers' extent of inbreeding, irrespective of whether this individual was actually their genetic father. Somewhat unexpectedly, there was no significant or strong effect from the mother's inbreeding. This indicates there may also be more indirect link between inbreeding and social behaviour acting across generations, whereby the male's inbreeding affects the environment of the chicks he is raising. While we did not directly measure parental care here, there is evidence that inbreeding causes parents to alter how they invest in their young, affecting resource allocation (Duthie et al., 2016) and the extent of care provided to young both before and after birth (Pooley et al., 2014; Wells et al., 2020). These conditions created by parents affect the overall raising environment experienced by their offspring, which has been shown to play a crucial role in determining the later life social strategy of juveniles via stress-linked effects (Boogert et al., 2014; Farine, Spencer, & Boogert, 2015), potentially through acting as a cue for environmental conditions (English et al., 2015). Previous studies in other passerines have also demonstrated that when juveniles experience particular environmental conditions also impacts on the magnitude and directionality of effects on their social behaviour (Boogert et al., 2013). Potentially, variation in parental roles between the sexes means that males and females contribute to care in different ways and at different times to exert differential effects

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on the raising conditions of their offspring (Buitron, 1988; McGhee & Bell, 2014). In hihi, while the female provides most care to nestlings, the male nevertheless contributes at this stage, and also becomes the primary care provider once chicks fledge. Therefore, the male has the potential to influence the rearing of his offspring over a prolonged period. Furthermore, the extent of male care is known to vary, at least in part due to paternity certainty (Ewen & Armstrong, 2000); this variation could therefore affect the rearing conditions of dependent offspring. As we cannot determine the exact mechanisms driving this result at this stage, obtaining measures of care across both nestling and fledging stages, alongside paternal inbreeding and EPC rates, would be an important next step to help explore the relationship between father's inbreeding and offspring sociality further. However, overall, this study provides the first evidence for consequences from indirect and intergenerational genetic effects on the sociability of young animals.

Our results support recent developments in understanding how complex social structure emerges from effects acting at a multitude of levels (Cantor et al., 2021), whereby the physical environment influence the likelihood of two animals ever forming an association (Strandburg-Peshkin et al., 2017), but individuals also exhibit their own social tendencies as a product of their traits and experiences which mediates associations once formed (Farine, Montiglio, & Spiegel, 2015). Here, we show that overall social structure in wild animals may emerge very early in life if the natal environment determines associations at both the individual and dyadic level. Furthermore, these patterns may even be pre-determined across generations if breeding and settlement decisions made by parents then determine the physical and social environments experienced by their offspring (Ilany & Akçay, 2016). For small populations in particular, the resulting social structure may have consequences for their evolutionary potential depending on individual mixing, particularly if associations continue to determine reproductive decisions across generations. For example, if early-life associations correlate with later breeding decisions, reflecting patterns found in other wild bird species such as great tits (Parus major; Firth & Sheldon, 2016), then this has the potential to impact on genetic mixing and inbreeding potential. However, such conditions may also promote evolution of alternate reproductive strategies in order to buffer from the risk of inbreeding: for example, in hihi we found that genetic relatedness was more independent of spatial proximity which may be a result of social strategies such as extra-pair mating or post-copulatory mechanisms of inbreeding avoidance (Brekke et al., 2012). Such consequences of early-life social structure remain to be tested explicitly, but are nevertheless crucial to understand as their potential implications highlight how early-life conditions may scale up to fundamentally determine population dynamics and evolutionary potential via the social environment.

### AUTHOR CONTRIBUTIONS

VF, PB and RT designed the study, VF led the data curation and analysis, with input from PB and RT; VF wrote the first draft of the manuscript; VF, PB and RT all contributed to manuscript revisions.

### ACKNOWLEDGEMENTS

We thank Matthew Silk for providing invaluable advice on initial social network analysis and Mhairi McCready and Donal Smith for monitoring the hihi population between 2015 and 2017. We are grateful to the Hihi Recovery Group, New Zealand Department of Conservation and Supporters of Tiritiri Matangi for their permission to conduct this research and supporting the long-term monitoring of the Tiritiri Matangi hihi population. We acknowledge Ngati Manuhiri as Mana Whenua and Kaitiaki of Te Hauturu-o-Toi and its taonga, remnant population and source of hihi. We also thank our reviewers for their helpful comments on the article. VF was supported by an Association for the Study of Animal Behaviour (ASAB) Research Grant and PB was supported by Research England.

### CONFLICT OF INTEREST STATEMENT

The authors declare no conflicts of interest.

### DATA AVAILABILITY STATEMENT

Data and R code used in this study are available via FigShare. doi: 10.6084/m9.figshare.22700116.v1 (Franks et al. 2023).

### BENEFIT-SHARING

All collaborators are included as co-authors in this study, and the results of research have been shared with all relevant parties and the broader scientific community. Benefits from this research accrue from the sharing of our data and results on public databases, as described earlier. Hihi are of cultural significance to the indigenous people of Aotearoa New Zealand, the Māori and are considered a taonga (treasured) species whose whakapapa (genealogy) is intricately tied to that of Māori. For this reason, the genotypes and associated pedigree for hihi will be made available by request on the recommendation of Ngāti Manuhiri, the iwi (tribe) that affiliates as kaitiaki (guardians) for hihi. To obtain contact details for the iwi, contact Dr Patricia Brekke.

### ORCID

Victoria R. Franks D https://orcid.org/0000-0003-4861-3036

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### SUPPORTING INFORMATION

Additional supporting information can be found online in the Supporting Information section at the end of this article.

How to cite this article: Franks, V. R., Thorogood, R., & Brekke, P. (2023). Parental breeding decisions and genetic quality predict social structure of independent offspring. *Molecular Ecology*, 00, 1–13. <u>https://doi.org/10.1111/</u> mec.17066