

STOCHASTIC PROJECTIONS FOR WATER RESOURCES PLANNING

A Report by

Leo R. Beard, Technical Director

and

Shin Chang, Research Associate

Center for Research in Water Resources

The University of Texas at Austin

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Foreword

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STOCHASTIC PROJECTIONS FOR WATER RESOURCES PLANNING

1. PROJECT OBJECTIVES

Evaluation of the future benefits and costs of water resources projections have traditionally been based on unique projections of key variables such as population and interest rates. Recognizing that the actual changes or trends that will occur in the future will be somewhat different from the theoretical projections, there is a fundamental need to evaluate the differences in benefits and costs that will occur if different trends occur. Also, the more comprehensive nature of water resources planning that is required today involves projections of many variables and, in projecting these variables, the interrelationships must be taken into consideration.

The objective of the research project reported on herein is to develop a technique for evaluating a range of projections of several types of variables that are involved in water resources planning and to develop techniques for using these projections effectively and efficiently. Variables used for illustration are economic, demographic and hydrologic. Only successive annual values are considered, since the project study is not concerned with seasonal variations. Only mathematical projection techniques are considered, since it is the technique of projecting sequences that is of concern to this project, and not developing cause-effect relationships wherein the projection of one variable can be used in deducing the projection of a second variable. Further, it is not an objective of this study to determine or develop best mathematical functions for projecting specific types of variables, but to suggest procedures, techniques and criteria for assessing the entire range of future potential as contrasted with the single most likely future.

2. BASIC CONSIDERATIONS

Of primary importance in most water resources planning studies, in

regard to projections of pertinent variables, is the projection level that will be reached at some specified time in the future. Of secondary importance are variations that occur along the line of projection. In most past studies, single projections of each pertinent variable have been made, and this ordinarily represents a smooth function that arrives at the most likely projection levels at various future times. Considering the uncertainty that available data on past quantities adequately represent future potential, and considering the virtual certainty that the actual future variations will not coincide with the most likely projections, water resources studies should consider alternative projection studies that can occur consistently with past variations of the particular variable being projected and considering our knowledge of how frequently and severely unexpected variations from indicated trends have occurred in the past.

As in the formulation of any mathematical model, the degree of complexity of a projection model must be adequate to represent variations of the phenomenon being studied with a degree of accuracy necessary for the particular applications anticipated. On the other hand, the model must be simple enough so that it can be calibrated satisfactorily with available data. Often it is not possible to meet both of these requirements, in which case the latter requirement would normally control, since there is little point in having a model that cannot be adequately calibrated. When this occurs, extreme caution must be used in interpreting the predictions that are made with the model.

Projection models described below have been restricted to a modest degree of complexity because of the limitations of data for calibration. Nevertheless, they include two types of variation that have not been used in traditional projection studies. These consist of a range of projections that adequately represent future possibilities and their associated probabilities, and variations within each projection that represent differences in trends that can be expected to occur from time to time. A small number of such projections covering the entire range of future expectations can be used separately to evaluate the accomplishments of proposed water resource projects, and expected project benefits can be computed by adding the cross products of the benefits obtained for each projection and the associated probability attached to that projection.

3. SINGLE-VARIABLE PROJECTION TECHNIQUES

Time series definitions. Some terms of special relevance to this report are defined as follows:

1. Time Series. A time series is a set of observations generated sequentially in time.
2. Stochastic Process. A stochastic process is a statistical phenomenon that evolves in time (or space) according to probabilistic laws. The time series to be analyzed may then be thought of as one particular realization, produced by the underlying probability mechanism, of the system under study.
3. Stationary Time Series. A time series is said to be stationary if the imbedded stochastic process of the time series remains in equilibrium about a constant level. Statistically, an nth-order stationary series requires that the first n moments of its distribution function be constant. For example, a second-order stationary series must have constant mean and constant variance.
4. Backward Shift Operator B. The backward shift operator B on a variable Z is defined by the equation: $B^m Z_t = Z_{t-m}$.
5. Backward Difference Operator V. The backward difference operator ∇ on a variable Z is defined in terms of B as: $\nabla = (1-B)$; hence $\nabla Z_t = Z_t - Z_{t-1}$.
6. Differences of a Time Series. Differences of a time series are obtained by successively subtracting the previous value from the current one, and this can be repeated to obtain differences of higher order. The degree of differencing is designated d. Using the backward difference operator ∇ , the dth difference of a time series Z can be represented as $\nabla^d Z$. For example, $\nabla^2 Z = \nabla(\nabla Z) = Z_t - 2Z_{t-1} + Z_{t-2}$.

Stationary Stochastic Models. Three categories of stationary stochastic models used in the description of stationary time series are autoregressive, moving average, and mixed autoregressive and moving average. These are defined as follows:

1. Autoregressive Models (AR). In these models, the current value

of the series is expressed as a finite, linear aggregate of the previous values plus a random component, a_t . Denoting the values of a time series at equally spaced times t , $t+1$, ..., by z_t, z_{t+1} , ..., and denoting the deviation from μ , the mean value of the time series, by \bar{z}_t, \bar{z}_{t+1} . Then the p -th AR model has the form

$$\begin{aligned}\mu &= (\sum_{t=1}^n z_t) / n, \\ \bar{z}_t &= \phi_1 \bar{z}_{t-1} + \phi_2 \bar{z}_{t-2} + \dots + \phi_p \bar{z}_{t-p} + a_t\end{aligned}\quad (3.1)$$

where $\phi_1, \phi_2, \dots, \phi_p$, are p 's autoregressive parameters. The model contains $p+2$ unknown parameters, $\mu, \phi_1, \dots, \phi_p$ and σ_a^2 , to be identified from the data. The additional parameter σ_a^2 is the variance of the white noise process or random component, a_t . The stochastic AR model can be useful in the representation of certain practical series. For instance, the first order and second order Markov models frequently used in streamflow synthesis are AR models.

2. Moving Average Models (MA). In these models, the current value of the series, \bar{z}_t , is expressed as a finite weighted sum of q previous random terms a_t . Thus the q -th MA model has the form

$$\bar{z}_t = a_t - \theta_1 a_{t-1} - \theta_2 a_{t-2} - \dots - \theta_q a_{t-q} \quad (3.2)$$

where $\theta_1, \theta_2, \dots, \theta_q$ are q 's moving average parameters. This model contains $q+2$ unknown parameters $\mu, \theta_1, \theta_2, \dots, \theta_q$ and σ_a^2 to be estimated from the data. Although this is called a moving average model, it is conceptually an infinite-term autoregressive model of specific form. The MA model is of practical importance in the representation of observed time series.

3. Mixed Autoregressive and Moving Average Models (ARMA). To achieve greater flexibility in fitting of actual time series, it is sometimes advantageous to include both autoregressive and moving average terms in one model. This leads to the mixed ARMA model. An ARMA(p, q) with p autoregressive terms and q moving average terms has the form,

$$\bar{Z}_t = \phi_1 \bar{Z}_{t-1} + \dots + \phi_p \bar{Z}_{t-p} + a_t - \phi_1 a_{t-1} - \dots - \phi_q a_{t-q} \quad (3.3)$$

The ARMA(p,q) model requires p+q+2 unknown parameters to be estimated from the data. It should be noted that the ARMA(p,q) model becomes an AR(p) model if q is set to zero and an MA(q) model if p is set to zero.

Nonstationary Stochastic Models. There are an unlimited number of ways in which a time series can be nonstationary. In fact, many time series in industry, business, and economics are better represented as nonstationary. Frequently, however, the nonstationary time series being analyzed exhibit a particular kind of homogenous nonstationary behavior. (i.e. the series exhibits homogeneity in the sense that, apart from the local level, or local level and trend, one segment of the series behaves much like any other segment). Differences of a nonstationary series of this kind can usually be expected to be stationary. Given the condition that the d-th difference of a nonstationary series is stationary, the autoregressive-integrated moving average model (which is a modified form of the ARMA model), can be used to describe the nonstationary series.

Autoregressive-Integrated Moving Average Models (ARIMA). The stochastic, nonstationary ARIMA model (with p autoregressive terms and q moving average terms) which describes a nonstationary series that is stationary in its dth difference has the form,

$$(1-\phi_1 B - \dots - \phi_p B^p) W_t = \theta_0 + (1-\theta_1 B - \dots - \theta_q B^q) a_t \quad (3.4)$$

with

$$W_t = \nabla^d Z_t, \text{ if } d > 0$$

$$= Z_t - \mu, \text{ if } d = 0$$

where Z is the series to be modeled,

W is the d th difference of the series Z ,

θ_0 is the long term trend factor.

It should be noted that the ARIMA(p,d,q) model becomes an AR(p) model if q and d are zero and an MA(q) model if p and d are zero and an ARMA model if d is zero.

Specific Models Studied. Seven models were selected for consideration under this project. They are

1. Second-Order Linear Markov Model.

$$z_t = a c_1 z_{t-1} + (1-a) c_2 z_{t-2} \quad (3.5)$$

where a is the weighting factor, c_1 and c_2 are random variables; $c_1 = z_t/z_{t-1} + b_1 t$ and $c_2 = z_t/z_{t-2} + b_2 t$ and b_1 and b_2 are stochastic linear regression coefficients, and t is time. The values of b_1 and b_2 for forecasting are estimated stochastically as a function of Student's t distribution from the known historical data.

2. Second-Order Nonlinear Markov Model.

$$z_t = z_{t-1}^{a c_1} \cdot z_{t-2}^{(1-a) c_2} \quad (3.6)$$

This model is the non-linear version of model 1. By taking logarithms of the variable, we have

$$\ln z_t = a c_1 \ln z_{t-1} + (1-a) c_2 \ln z_{t-2}, \quad (3.7)$$

this equation has the same form as that of model 1 except that the logarithms of the time series Z are used instead of the time series itself. As in model 1, c_1 and c_2 are estimated stochastically.

3. Nth-Order Linear Markov Model.

$$Z_t = aC_1 Z_{t-1} + bC_2 Z_{t-2} + cC_3 Z_{t-3} + dC_4 Z_{t-4} + eC_5 Z_{t-5} \text{ for } N \leq 5. \quad (3.8)$$

The factors a, b, c, d, and e provide for weighting of the independent variables. The linear coefficients C_1 , C_2 , C_3 , C_4 , and C_5 are random variables and are distributed according to certain probability functions. For the purposes of this study, these coefficients were assumed to be normally distributed and defined by the equation

$$C_i = Z_t / Z_{t-i}, \quad i = 1, 2, 3, 4, 5 \quad (3.9)$$

When the model is being used for forecasting, the values of the linear coefficients are estimated stochastically from the historical data.

4. Nth-Order Nonlinear Markov Model.

$$Z_t = Z_{t-1}^{aC_1} \cdot Z_{t-2}^{bC_2} \cdot Z_{t-3}^{cC_3} \cdot Z_{t-4}^{dC_4} \cdot Z_{t-5}^{eC_5} \quad (3.10)$$

This model is the nonlinear version of model 3. By taking logarithms of the variable, the resulting equation has the same form as that of model 3.

5. Decade and Annual Ratio Autocorrelation Model. This model takes advantage of the availability of 10-year census data and the relatively accurate estimates of annual data that have been obtained using census techniques and omitting annual data obtained by mathematical interpolation. This permits the most effective use of early census information when only decade values are available. The model consists of a lag-1 Markov chain of the logarithms of the ratios of successive decade values and an interpolation routine using a similar Markov chain for annual logarithms where annual observations are available. When generating successive sequences, random values of average decade ratios are used for each sequence rather than sample values.

6. Decade Regression and Annual Autocorrelation Model. This model is similar to the preceding model except that a linear regression on the logarithms of the decade ratios is used instead of autocorrelation, and a new random value of the regression coefficient is used for each generated sequence.

7. The ARIMA(p,d,q). The ARIMA model includes three classes of models previously described: the AR, the MA, and the ARMA. It can be applied to both stationary and nonstationary time series. The degree of differencing, d, and the number of parameters, p and q, are identified first, then values of parameters are estimated by the least mean square errors method.

The first six models are essentially special cases of the AR process and they all experienced forecasting problems in some applications, particularly in the nature of forecasting unreasonable quantities. Thus, the ARIMA model was chosen as the standard forecasting model because of its flexibility and adaptability to various types of time series, thus providing means of avoiding unreasonable forecasts.

4. THE ARIMA MODEL

ARIMA(p,d, q) Model Building. To fit a series by an ARIMA(p,d,q) model, the model building process is ordinarily based on available observations. Computationally, the model building includes three steps: identification, estimation, and diagnostic checking. The order of differencing, d, the number of autoregressive parameters, p, and the number of moving average parameters, q, are identified in the identification step; values of parameters are then estimated in the estimation step; and the adequacy of the model in describing and projecting the series is checked in the diagnostic step.

According to Box and Jenkins (1970), no pure mathematical approach can be adopted in the model building process. Graphical methods and judgment are needed throughout because of the trial-and-error nature of the process. Following their basic ideas, a systematic noniterative algorithm is developed for model building, and although the algorithm is conceptually an exhaustive trial-and-error approach, operationally it is automatic. All possible models from the general ARIMA family of limited

order are examined, and the best model is selected such that four proposed performance criteria are either satisfied or maximized. These performance criteria are established to measure the goodness of fit and reasonableness of projections.

1. Identification of the Order of Differencing d. The failure of the autocorrelation function to die out quickly is taken as an indication that the underlying stochastic process of the series Z is nonstationary, but the process of its first or higher order differences may be stationary. Therefore, it is assumed that the degree of differencing, d , necessary to achieve stationarity has been reached when the autocorrelation function of $W_t = \nabla^d Z_t$ dies out quickly. In practice, d is generally 0, 1, or 2, and it is usually sufficient to inspect the first 20 (or so) estimated autocorrelations of the original series and of its first and second differences. According to Box and Jenkins (1970), to judge whether or not the autocorrelation function dies out quickly is completely subjective, as no standard judging criterion is available. The following method was found to be an effective way of detecting the speed of decay of the autocorrelation function and identifying the order of differencing, d .

Given a series of n elements and denoting the lag i autocorrelation coefficient by r_i , the Chi-square statistic, χ_d^2 , of the first 24 autocorrelation coefficients can be calculated as

$$\chi_d^2 = n \cdot \sum_{i=1}^{24} r_i^2 , \quad (4.1)$$

Where d can be 0, 1, or 2 to stand for, respectively, the Chi-square statistic of the original series, of its first difference, or of its second difference. Let R_1 and R_2 be the ratios of

$$R_1 = \chi_1^2 / \chi_0^2, \text{ and } R_2 = \chi_2^2 / \chi_1^2 \quad (4.2)$$

In general, R_1 and R_2 are less than one. These ratios give the relative decay speed of the autocorrelation function of the series and its differences. For example, if $R_1 = 0.2$, then the Chi-square statistic of the

first difference is only 0.2 times that of the original series. This indicates that the speed of decay of the autocorrelation function of the first difference is five times faster than that of the original series. Based on this argument, the following rule can be used to select the degree of differencing d for the model:

$$\begin{aligned}\text{Select } d &= 0, \text{ if } R_1 > \epsilon \text{ and } R_2 > \epsilon; \\ d &= 1, \text{ if } R_1 < \epsilon \\ d &= 2, \text{ if } R_1 > \epsilon \text{ and } R_2 < \epsilon.\end{aligned}$$

The criterion, ϵ , is set from experience and judgment. A value of 0.4 was used in identifying d for all sample series in this study.

Parameter Estimation. Once it has been determined what the degree of differencing should be, values of parameters in all the member models of the general ARIMA family are next estimated. In the general order of decreasing complexity, the fourteen most used members of the ARIMA(p,d,q) family (listed by their (p,d,q) values) are: $(2,d,3)$, $(3,d,2)$, $(1,d,3)$, $(2,d,2)$, $(3,d,1)$, $(0,d,3)$, $(1,d,2)$, $(2,d,1)$, $(3,d,0)$, $(0,d,2)$, $(1,d,1)$, $(2,d,0)$, $(0,d,1)$, $(1,d,0)$. Box and Jenkins suggest that several of the models most likely to best describe the time series can be identified by comparing the shapes of their autocorrelation and partial autocorrelation functions with those of certain established theoretical patterns. The best model is then selected by a detailed study of the residuals. For the following reasons, however, it is believed that the following approach is more attractive and easier to apply than that of Box and Jenkins:

1. From the experience of modeling 30 sample series, it was found that in most cases, the autocorrelation and partial autocorrelation functions of the appropriately differenced series fail to show any pattern to yield a meaningful conclusion in selecting p and q in accordance with the findings of Box and Jenkins.

2. In some cases, models with different values of p and q can be fitted equally well to a given series and yield equally high levels of significance in the diagnostic checking step. This in turn suggests that

correspondence between the selection of p and q of an ARIMA(p,d,q) model and the shape of autocorrelation and partial autocorrelation functions is not unique.

3. There is no definite trend in either direction between the complexity (number of parameters used in a model) of a model and the performance (to be discussed in the model screening step) of the model. In other words, no selective search technique, rather than an exhaustive search of all possible models, can be developed to find the best model.

4. No plot of the autocorrelation and the partial autocorrelation function is required for the selected approach.

5. No intermediate subjective feedback is needed in the course of selecting the best model by the current approach; hence an automatic computer search procedure is possible.

The estimation of autoregressive and moving average parameters for an ARIMA(p,d,q) model is based on the Marquardt algorithm for nonlinear least-square parameter estimation. Computer programs based on this algorithm are available from the IMSL (International Mathematical and Statistical Libraries) program library 3 (for CDC Cyber 70/6000/7000), in which FTARPS and FTMAPS (both in FORTRAN) were used in this study to estimate the autoregressive and moving average parameters, respectively. After estimating the parameters of an ARIMA(p,d,q) model, the residuals, a_t , are computed by the equation

$$a_t = \hat{W}_t - W_t, \quad (4.3)$$

where \hat{W}_t is the fitted value of the time series and W_t the observed value.

Model Screening. Four criteria for selecting the best model to describe a time series are used:

1. Goodness of Fit. The goodness of fit of an ARIMA model to a time series is measured by the level of significance of the Chi-square test for the residual, a_t , in relation to white noise (randomness). The Chi-square statistic of the residual a_t for $t = 1, 2, \dots, n-d$, is calculated using the formula,

$$\chi^2 = (n-d) \sum_{i=1}^m r_i^2 \quad (4.4)$$

where r_i is the i th autocorrelation coefficient of series a_t , and m is the number of autocorrelation coefficients used in the computation of the statistic. With $m-p-q$ degrees of freedom, the significance level associated with χ^2 can be found from a Chi-square test table. The computer program MDCH of IMSL was used to find this level. The resulting residuals, a_t , should not significantly differ from a white noise process, if the time series of interest is fitted by an adequate ARIMA model. Thus, a minimum level of significance can be set to reject inadequate models. A minimum level of 0.01 was used in this study.

2. Model Stability. There exists an admissible region of parameters for every ARIMA(p,d,q) model which will ensure the stationarity and invertibility of the model (see chapter 3 of Box and Jenkins, 1970). Models with parameters outside this region will be dropped from further consideration. The region can be expressed implicitly as:

a. For Autoregressive Parameters. The roots of the polynomial $(1-\phi_1B-\phi_2B^2 - \dots - \phi_pB^p)$ which is formed by taking all autoregressive terms of the ARIMA(p,d,q) model (i.e. left hand side of equation 3.4) when solved for B , must be greater than one.

b. For Moving Average Parameters. The roots of the polynomial $(1-\theta_1B-\theta_2B^2 - \dots - \theta_qB^q)$, which is formed by taking all moving average terms of the ARIMA(p,d,q) model (i.e. right hand side of equation 3.4), when solved for B , must be greater than one.

3. Reasonableness of Forecasting. A model that can be fitted to a series adequately in the sense of passing the minimum level of significance in the goodness of fit test might nevertheless give unreasonable future projections, such as interest rates of 1000 percent. This is especially true when split-record tests are performed on a U-shaped or a W-shaped series. Therefore, a well-fitted model with a high level of significance is not necessarily a good projection model for the purpose of planning. A projection-reasonableness test is used as a supplementary criterion to the level of significance in selecting a model. A physical range and/or lower and upper bounds for the future are established based on available information and basic characteristics of the series. In this study, the bounds

are established at a distance of 50 times the range of values within the record from each extreme of record in terms of the transformed variable (logarithm, usually). A model is said to have projection reasonableness with respect to a series if projections of the series made by this model can satisfy all the designated range limitations. In this test, five independent projection sets, each with five times the record length, are generated randomly by the model. A model which cannot pass this test will be discarded.

4. Simplicity of Model. According to the principle of parsimony in using parameters, one should not try to overfit a time series with a complicated model when a simpler model can fit the series adequately. Thus, a trade-off between the performance (measured by the level of significance in the goodness of fit test) and the complexity (measured by the total number of parameters used) of models is desirable in selecting the best model for practical uses. The trade-off can be achieved by employing a discount factor, i , to the level of significance, α , such that α is reduced exponentially according to the complexity of the model. Mathematically, let $(p+q)$ be the degree of complexity of an ARIMA (p,d,q) , then the discounted level of significance, α' , is computed as

$$\alpha' = \alpha(1-i)^{p+q-1} \quad (4.9)$$

In selecting a value of i , personal judgment must be exercised. A value of 0.3 was used in this study, and this represents a high degree of discrimination against the more complex models.

Following the above four model screening steps, the best model can be selected simply by choosing the model which has the highest discounted significance level.

Applications of the ARIMA(p,d,q) Model. The ARIMA(p,d,a) model can be used to fit, to forecast, and to simulate a given time series.

1. Fitting a time series. Let \hat{W}_t be the fitted value of the observed time series W_t and a_t be the residual of fitting \hat{W}_t to W_t at time t . The first fitted value of \hat{W}_t to be computed using the fitted ARIMA model is \hat{W}_{p+1} , because the first p values of W_t are not used in the estimation

of autoregressive parameter ϕ_p . Thus for $t \leq p$, we have

$$\hat{W}_t = W_t \quad (4.5)$$

$$a_t = 0 \quad (4.6)$$

and for $t > p$, we have

$$\hat{W}_t = \phi_1 \hat{W}_{t-1} + \dots + \phi_p \hat{W}_{t-p} + \theta_0 - \theta_1 a_{t-1} - \dots - \theta_q a_{t-q} \quad (4.7)$$

$$a_t = \hat{W}_t - W_t, \quad (4.8)$$

and in equation (4.7), the random noise term, a_j with negative lag j is set to zero. It should be pointed out that W is the d th difference of the original observed series Z . Let \hat{Z}_t be the fitted value of the observed Z_t at time t , then the functional relationship of Z_t can be derived by substituting $W_t = \nabla^d Z_t$ into equation (4.7). Upon substituting and collecting similar terms, we have

$$\hat{Z}_t = \sum_{i=1}^{p+d} \psi_i \hat{Z}_{t-i} + \theta_0 - \sum_{i=1}^q \theta_i a_{t-i} \quad (4.9)$$

where ψ_i 's are coefficients of \hat{Z} . For example, if W is the first difference of Z_t (i.e. $W_t = \nabla Z_t = Z_t - Z_{t-1}$), then, by substituting $\hat{W}_t = \hat{Z}_t - \hat{Z}_{t-1}$ into equation (4.7), we have

$$\begin{aligned} \hat{Z}_t - \hat{Z}_{t-1} = & \phi_1 (\hat{Z}_{t-1} - \hat{Z}_{t-2}) + \phi_2 (\hat{Z}_{t-2} - \hat{Z}_{t-3}) + \dots + \\ & \phi_p (\hat{Z}_{t-p} - \hat{Z}_{t-p-1}) + \theta_0 - \theta_1 a_{t-1} - \dots - \theta_q a_{t-q}, \end{aligned} \quad (4.10)$$

collecting similar terms,

$$\begin{aligned} \hat{Z}_t = & (1-\phi_1) \hat{Z}_{t-1} + (\phi_2-\phi_1) \hat{Z}_{t-2} + \dots + (-\phi_p) \\ & \hat{Z}_{t-p-1} + \theta_0 - \theta_1 a_{t-1} - \dots - \theta_q a_{t-q} \end{aligned} \quad (4.11)$$

or, in terms of ψ_i ,

$$\hat{Z}_t = \sum_{i=1}^{p+1} \psi_i \hat{Z}_{t-i} + \theta_0 - \sum_{i=1}^q \theta_i a_{t-i} \quad (4.12)$$

2. Forecasting and Simulation. Let \hat{W}_{t+s} be the forecast of the series W_t at lead time s from the origin t . It can be shown (Box and Jenkins, 1970, Chapter 5) that: (a) the forecast for W_{t+s} using minimum mean square error estimation is the conditional expectation of W at time $t+s$; (b) the expected forecast error of \hat{W}_{t+s} is equal to zero, hence the forecast is unbiased; and (c) the one-step-ahead forecast error is equal to a_t (the residual at time t), hence the Chi-square significance test of a_t can be used as an effective indicator of the goodness of fit of the ARIMA model for the series being modeled. The forecasting equation for \hat{W}_{t+s} is

$$\hat{W}_{t+s} = \phi_1 \hat{W}_{t+s-1} + \dots + \phi_p \hat{W}_{t+s-p} + \theta_0 - \theta_1 a_{t+s-1} - \dots - \theta_q a_{t+s-q} \quad (4.13)$$

where $\hat{W}_t = \begin{cases} W_t, & \text{if } t \leq n-d, \\ \hat{a}_t, & \text{if } t > n-d, \end{cases}$
 $a_t = \begin{cases} a_t, & \text{if } t \leq n-d, \\ E(a_t) = 0, & \text{if } t > n-d \text{ and } E(a_t) \text{ is the expected value} \\ & \text{of the white noise } a_t. \end{cases}$

And similar to the derivation of the fitted value \hat{Z}_t from \hat{W}_t , the forecasting function for \hat{Z}_{t+s} can be shown to be

$$\begin{aligned} \hat{Z}_{t+s} = & \psi_1 \hat{Z}_{t+s-1} + \dots + \psi_{p+d} \hat{Z}_{t+s-p-d} + \theta_0 - \theta_1 a_{t+s-1} \\ & - \dots - \theta_q a_{t+s-q} \end{aligned} \quad (4.14)$$

where $\hat{Z}_t = \begin{cases} Z_t, & \text{if } t \leq n \\ \hat{a}_t, & \text{if } t > n, \text{ and} \end{cases}$
 $a_t = \begin{cases} a_t, & \text{if } t \leq n \\ 0, & \text{if } t > n. \end{cases}$

The forecast \hat{Z}_{t+s} represents the most likely projection level for lead time s in the future. For reasons mentioned before, alternative projections are needed and can be generated by simulation using the fitted ARIMA model with random components.

Each independent set of simulated projections is generated using a

stochastic trend factor β and a white noise term b_t . The stochastic trend factor β is obtained by assuming that the trend factor, θ_0 , used in forecasting, has a Student-t distribution. The white noise term b_t is generated by a random variable generator using the known variance of the white noise a_t . That is

$$\begin{aligned} b_t &= a_t \text{ if } t \leq n \\ &= q \sqrt{\sigma_a^2}, \text{ if } t > n \end{aligned} \quad (4.15)$$

q is a normally distributed variable with zero mean and unit variance. Then, the simulation function of \hat{Z}_{t+s} is

$$\begin{aligned} \hat{Z}_{t+s} &= \psi_1 Z_{t+s-1} + \dots + \psi_{p+d} Z_{t+s-p-d} + \beta - \theta_1 b_{t+s-1} \\ &\quad - \dots - \theta_q b_{t+s-q}, \end{aligned} \quad (4.16)$$

$$\begin{aligned} \text{where } \hat{Z}_t &= Z_t \text{ if } t \leq n \\ &= Z_t \text{ if } t > n. \end{aligned}$$

5. MULTIVARIATE PROJECTION MODEL

General Approach. For variables that are interrelated, an acceptable model should project every variable simultaneously so as to preserve the interrelationship among projections of all variables. There is no specific multivariate projection technique, but principal-component analysis can be used where data for all variables exist for the same period, in order to form uncorrelated variables (principal components), each of which is expressed as a linear function of the original variables. Then the new variables can be projected independently and transformed to form projections of the original variables. The principal-component technique was selected for convenience in this study as an illustrative procedure. Other multivariate techniques, such as multiple linear regression, can be used if desired.

Principal-Component Analysis. Principal-Component analysis is a straight-forward method of transforming a given set of variables into a new set of composite variables (or principal components) that are mutually orthogonal (uncorrelated). The first principal component represents the single best summary of linear relationships exhibited in the data. The second component is defined as the second best linear combination of variables that accounts for the most residual variance after the effect of the first component is removed from the data. Subsequent components are defined similarly until all the variance in the data is exhausted. Unless at least one variable is perfectly determined by the remaining variables in the data, the principal-component solution requires as many components as there are variables. Let Z_j , $j=1, 2, \dots, m$ be the m original variables; F_k , $k=1, 2, \dots, m$ be the m principal components; A_{jk} , $j=1, 2, \dots, m$ and $k=1, 2, \dots, m$ be the component-pattern coefficients, and b_{jk} , $j=1, 2, \dots, m$ and $k=1, 2, \dots, m$ be the component-estimation coefficients. Then the principal component F_k can be computed from the values of the m original variables Z as

$$F_k = b_{k1} Z_1 + b_{k2} Z_2 + \dots + b_{km} Z_m, \quad k = 1, 2, \dots, m \quad (5.1)$$

and the original variable Z_j can be described by the m principal components F as

$$Z_j = A_{j1} F_1 + A_{j2} F_2 + \dots + A_{jm} F_m, \quad j = 1, 2, \dots, m. \quad (5.2)$$

It should be noted that a reduction in the number of principal components used in equation 5.2 is possible if a certain percentage of unexplained variance of the m original variables Z is allowed. For example, if the first m_1 ($m_1 < m$) principal components account for 98 percent of the variance (hence 2 percent unexplained variance) in the m original variables, and if a 2 percent unexplained variance in expressing the m original variables is allowed, then only the first m_1 principal components are needed in describing the m original variables. The value of m_1 is dependent on

the value of m and the maximum percentage of unexplained variance of the original variables allowed. The reduction in the number of principal components used from m to m_1 represents a way of reducing the size of the multivariate projection model, since only projections of m_1 principal components are needed to calculate projections of m original variables. In the demonstration examples herein, no reduction in variance was allowed, hence $m_1 = m$ in these cases.

Multivariate Projection Model. Simultaneous projection of a set of original variables can be obtained by the following steps using the principal-component analysis allied with the single-variable projection model, ARIMA(p, d, q):

1. Find the component-pattern coefficients, A_{jk} , and component-estimate coefficients, b_{jk} , by principal-component analysis. (For the purpose of this study, these coefficients were calculated by the computer program BMD01M which is written in FORTRAN and available from Biomedical Computer Programs (BMD) developed at UCLA for IBM system 360/370.)
2. Construct a time series for each of m principal components by substituting values of the m original variables in equation 5.1.
3. Apply the ARIMA(p, d, q) model as described in section 4 to project future values of the most likely projection level and of a number of simulation sets for each of m principal components independently.
4. Compute projections of the most likely projection level and simulations of the m original variables by substituting projections of principal components in equation 5.2.

6. TEST TECHNIQUE

Split-Record Testing. In order to develop some degree of confidence that projections using techniques studied herein would in fact represent the entire range of future possibilities, or at least to evaluate the degree to which they do represent the future range and possibly to develop correctional criteria, the split-record type of testing has been employed in this study. This consists of calibrating the projection models on the basis of one part of the record (usually half) and examining the range of projections in comparison with the actual variations that occurred in the remainder of the record. This technique has been particularly favored, because it conforms to the exact manner in which projections are employed in water resources planning. In the test procedure, the first half of the record might represent all of the record that would be available at the time of planning the project, and the second half of the record would represent what occurs during the operation of the project. Thus, in concept, the test would show how the various projections made in project planning would actually compare to events that really occurred during the project life.

The split-record technique should be evaluated on the basis of a substantial number of variables. It is possible that a particular technique might accidentally show exceptional results when applied to one or two variables, but it is not likely that consistently superior results will be accidentally obtained as the number of variables increases. Even when a substantial number of test variables are used, care must be exercised that a great many techniques are not arbitrarily tested and the best selected, because, if sufficient alternatives are tried, there is an increasing probability that one will accidentally show favorable results.

When records are studied in halves for the purpose of testing, it is often possible to make more efficient use of available data by reversing the halves and repeating the test procedure. Thus, a model calibrated on the first half could be used for projecting through the period of the second half, and a model calibrated in reverse on the second half could be used for projecting in reverse through the first half of the record. Depending on the nature of the trends, this double use of the data could actually

double the effectiveness of the test, but in cases where the trends are consistently upward or downward, this double use of the data might not contribute greatly to the test results insofar as they would apply to those particular variables. In order to maximize the similarity of the tests to actual application, halves were not reversed in this study.

It should be recognized that, even though the test procedures are limited to calibration of the model using only half of the data, actual projections for water resources planning studies should be based on models calibrated using all available data.

Sample Time Series. To facilitate the test, 31 time series of 17 economic, 10 population, and 4 hydrologic variables were collected and listed in table 1.

Selection of Representative Projections. In each split-record test, 51 stochastic projections were made. These projections are numbered in the order of magnitude of the final value at the end of the projection (end of the second half of the record). In order to illustrate the range of projections and, in fact, to represent the entire set of projections with a small number of projections, projection numbers 2, 8, 26, 44 and 50 were selected. These are illustrated in figures 1 to 37. In each figure, the heavy line represents the observed data, the line connecting the symbol, \oplus , represents the most likely projection level, and the other five plots represent the five selected representative projections. On the basis of the variables used in this study, it appears that variations can occur in the future with a substantial probability (in the order of once in 50 projections) that differ in magnitude from the extreme projections computed in this study by an amount exceeding the range of projections computed and illustrated herein. In order to take these "maverick" possibilities into account properly, it is proposed that the following weights be assigned to the five representative projections:

<u>Projection no.</u>	<u>Representing Projection</u>	<u>Probability</u>
2	1 - 3	.06
8	4 - 12	.18
26	13 - 39	.50
44	40 - 48	.18
50	49 - 51	.06
Total		<u>.98</u>

Two more projections, to be assigned .01 probability each, can be obtained by utilizing projections numbers 2 and 50 to obtain extreme projections A and B such that values for each year are defined as:

$$Z_A = 2Z_2 - Z_{50}$$

$$Z_B = Z_2 - 2Z_{50}$$

This gives 2 new (maverick) projections that differ from projections 2 and 50 by the difference between 2 and 50 values for each respective year in terms of the transformed variable, Z. The maverick projections would each have a weight (probability) of .01, thus completing the range of probability represented in the above tabulation. Figures 38 and 39 are used to demonstrate the use of these extreme projections for the variable Man-Hours in Farming and variable Construction Cost, respectively. In each chart, projection A is represented by a single line connecting the symbol, *, and projection B by a single bold line.

7. STUDY RESULTS

Analysis of Economic Variables. Of the 17 economic variables (Fig 1 to Fig 17), observed quantities at the end of the projection period were outside of the range of projection in 5 cases (Figures 6, 7, 8, 13, and 17). If the projection techniques adequately represented all future probabilities, only one of these would be expected to be outside the range (an expected frequency of 4 out of 52 random projections would lie outside the range of the second-largest and second-smallest of 51 projections).

A careful review of these projections will indicate that, while different types of models would have greatly improved the projections in individual cases, no systematic model will satisfactorily represent what actually happened in all cases. Likewise, once a person sees the outcome of an observed variable, it is often possible to explain the variations, usually in terms of related factors, but it should be remembered that this study is concerned only with statistical means of projecting a stochastic variable or a set of interrelated stochastic variables.

It becomes obvious, therefore, as has been generally recognized, that the probabilistic range of projections should include provision for possible radical changes not indicated by the model calibration data. This is partly due to the fact that no model simple enough to be calibrated satisfactorily can adequately represent the complex factors involved and partly due to the likelihood that factors of potentially major impact are not adequately reflected in a particular sample of data.

Analysis of Demographic Variables. Population projections were made individually for all 10 cases (Fig 18 to Fig 27), and then 6 cases (Fig 28 to Fig 33) were projected jointly. Models selected by the computer program are as follows:

<u>Principal-component No.</u>	<u>ARIMA(p,d,q) model</u>
1	0, 1, 1
2	1, 1, 0
3	1, 1, 0
4	0, 1, 1
5	2, 1, 0
6	0, 1, 1

In the joint projections, selection of extreme projections was based on the weighted total population at the end of the sequence, giving equal weight to each variable, and simultaneous projections for all variables were used. Thus, the second highest of 51 projections for the weighted total population was not necessarily the second highest projection for each variable. In fact, it would tend to be less extreme in most or all cases, thus reducing the range of projections for each variable below that obtained in the independent projections.

Whether the projections should be made jointly or independently would depend on the application, and it is just as vital that they be made independently for independent applications as it is that they be made jointly for applications where their simultaneous impact is of primary concern.

Of the 10 independent projections, observed values at the end of the projection period lie outside the projection range in 2 cases, which is at least twice the expected frequency. Thus, again it is necessary to account for greater variation in potential projections than a reasonably simple model indicates.

Analysis of Streamflow Variables. In the cases of natural streamflows, projections (Fig 34 to Fig 37) appear to be reasonable in relation to

observed variations during the period of projection, except for the Columbia River, where much lower quantities occurred in some years than were projected to any year. There is a strong downward trend through the early years and until 1935, after which a period of 20 years showed a strong resurgence of streamflows. The model used for streamflows is based on steady-state projection, because this gives best results for streamflows in general. If the model using $d = 1$ (first differences projected were used), it is apparent that the flows from 1935 to 1955 could fall well above all projections. It is difficult to determine whether the selected model is inapplicable or the observed data represent rare occurrences. There is a tendency to jump to the conclusion that actual events that conform to a forecasted projection "prove" the projection to be acceptable and that when actual events do not conform, the projection was in error. This, of course, is an over-simplification.

It is pertinent to recognize that the streamflow model determined to be most suitable is a steady-state model ($d=0$), and that the more general and extensively tested monthly streamflow simulation model developed in the Hydrologic Engineering Center of the Corps of Engineers is more suited for generation of sequences of this type of variable.

8. APPLICATION

General. In discussing the application of techniques studied under this project, it must be remembered that the techniques are meant to be applied to basic variables that cannot be derived directly from other variables that are more easily predictable. That is, the extrapolation is strictly statistical, as it must be in the case of many variables.

In making such extrapolations, natural or physical bounds must be recognized. The natural lower asymptotic bound of zero is taken into account in this study by transforming the variable to a logarithm, which is then unbounded at the lower end. If a variable is asymptotically bounded only on the upper end, taking the logarithm of the difference between the upper limit and the variable would remove that bound. If there are both lower and upper asymptotic bounds, the following transform can be used:

$$Z = \log\left(\log \frac{B-A}{X-A}\right)$$

where B is the upper bound, A the lower bound, X is the observed variable, and Z is the transformed variable. Other transforms might be devised and perform better under specific circumstances, but this one will at least assure that impossible values are not predicted. The transformed variable is projected, and resulting values are transformed back to the original variable needed in the contemplated study.

When more than one variable are needed in the study and where these are interdependent, it is necessary to perform a principal-component analysis of the transformed variables and to project these components, then recombine them into the transformed variables before transforming back to the original variables.

All of the above provisions are incorporated into the computer program described herein.

Water Resource System Simulation. Because of the complexity of water resources planning and management studies, the most acceptable general procedure for analysis consists of postulating a system and operation rules, and simulating the operation of the system under anticipated system demands and system inputs. Traditionally, system demands have been projected as a most likely sequence of

demands, and system inputs have been assumed similar to past experience, adjusted for future changes in controlling conditions (such as urbanization and reservoir construction). It is the thesis of this study that a range of demands and inputs should be considered and that appropriate weights should be assigned to the results of different projections. The technique described above for computing 51 projections and selecting 5 of these plus 2 "maverick" projections is intended to represent the full range of future probabilities with a minimum of computational effort.

Computer programs such as HEC-3, "Reservoir System Analysis for Conservation", developed in the Hydrologic Engineering Center of the Corps of Engineers, are very useful in performing the detailed simulation studies necessary for evaluating system performance under each combination of projections of the pertinent input and demand variables. In the case of hydrologic variables such as streamflow, rainfall and evaporation, where the process can be treated as stationary and where stochastic variations from month to month are important, they should be projected using a model such as HEC-4, "Monthly Streamflow Simulation".

Integration of simulation outputs. To the extent that any of the input and demand variables are independent of each other, it will be necessary to give special consideration to the manner in which the different sequences of the various variables are combined. In extreme cases, it may be necessary to perform simulations corresponding to all combinations of sequences, but it is hoped that this will not be necessary in ordinary applications. For example, principal variations in physical results of a plan of operation may depend almost entirely on hydrologic sequences, in which case the detailed water resource simulation studies need be made only once for each hydrologic sequence. Evaluation of the results in terms of economic output would require integration of many combinations of these outputs with the various economic projections. The expected system output would then be the sum of the cross products of the output for each simulation times the probability associated with the particular input. The probability associated with any particular input when a combination of independent inputs

is used equals the product of the probabilities associated with the independent inputs. It must be assured that probabilities associated with all simulations add to 1.0.

Illustrative Projections. For illustration purposes, projections of sample variables to year 2020 were obtained by using the projection techniques and computer programs developed in this study. These are illustrated in figures 40 to 61. Variables which are considered to be interdependent are projected by use of the multivariate projection model. These are:

1. Per capita GNP, industrial production index, gross nonfarm product, and gross farm product (shown in figures 40, 41, 42, 43, respectively).
2. Output per unit of labor input index, and output per unit of capital input index (shown in figures 44, 45 respectively)
3. Wholesale price index and consumer price index (shown in figures 46 and 47 respectively).
4. Manhattan Island real estate mortgage rates, stock exchange call loan rates and commercial paper rates (shown in figures 48, 49, and 50, respectively).
5. Population of New England, Middle Atlantic, East North Central, West North Central, South Atlantic, and East South Central divisions (shown in figures 51, 52, 53, 54, 55, 56 respectively).

Variables projected by single-variable projection models are man-hours in manufacturing, man-hours in farming, number of operating business, unemployment rate, applications for patents, and construction cost index, and projections are shown in figures 57 to 62, respectively. It should be noted that projections illustrated herein are used only to demonstrate the projection techniques developed in this study and should not be used as a formal projection of those variables.

Computer Programs. Algorithms described herein are combined for convenient use into two computer programs -- one for single-variable projections, described in Appendix A, and one for multivariate projections, described in Appendix B.

9. CONCLUSION

Comparison of actual sequences of demographic and economic variables demonstrates that no projection technique can be depended upon to yield acceptable results if only a single sequence is projected. It is not possible to predict variations accurately, so it is essential that a range of projections be considered if errors of projection are important to a proposed study.

A model used for projecting such variables must be flexible enough to define variations reasonably, yet simple enough to be calibrated satisfactorily with available data. The two requirements are often unobtainable, and it becomes necessary in those cases to conform with the latter requirement. As a consequence, estimated errors of projection are usually smaller than true errors, and it has become necessary to expand the range of possible projections as described herein.

The autoregressive-moving-average family of projection models described herein provides a range of models from which one can be selected on the basis of goodness of fit to observed data and reasonableness of projection as described herein. The selected model can then be calibrated on the basis of all available data, and a set of projections generated to constitute input and demand data for any water resource system simulation study. Computer programs for accomplishing this are described in appendices A and B. The weighted average accomplishments of any proposed plan of development or management can then be computed as described herein, for the purpose of evaluating the plan under the complete range of future possible projections and computing and evaluation that should be far superior to any obtainable by conventional methodology based on single projections of each variable.

10. REFERENCES

A search of recent statistical literature concerning time series analysis especially in the area of forecasting and control was undertaken in order to compare methodologies and to apply newly advanced treatments to the problems addressed by this study. It was found that most of the related material published prior to 1970 was included in the book entitled Time Series Analysis; Forecasting and Control (Box and Jenkins, 1970). Chatfield and Pepper (1971) discussed some practical aspects of time-series analysis, Bhattacharyya (1974) employed the forecasting technique in time-series analysis to predict the demand for telephones in Australia, and Box and Jenkins (1974) presented some of the more recent advances in forecasting and control of a time-series.

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TABLE 1
RECORDED ECONOMIC, DEMOGRAPHIC AND HYDROLOGIC DATA

	A11 PER CAPITA GNP (1958 DOLLARS, 1889=1966)	78	944
805	847	867	932
1913	1922	1119	1107
1259	1246	1266	1305
1446	1403	1344	1404
1748	1565	1442	1234
1699	1818	2086	2348
2367	2519	2655	2679
2899	2913	2917	3052
	A15 INDUSTRIAL PRODUCTION INDEX (1913=100, 1860=1959)	100	100
7.48	7.49	6.94	7.86
11.78	12.38	14.60	14.40
20.58	22.38	23.90	24.40
35.08	36.08	38.80	34.70
58.68	56.78	63.28	65.40
85.38	82.28	93.78	90.00
124.08	100.18	125.90	144.40
155.68	129.78	160.50	119.90
213.98	275.58	340.38	405.20
366.38	398.78	411.70	447.30
	A17 GROSS NONFARM PRODUCT (BILLIONS OF DOLLARS) 1889=1966)	78	944
33.4	37.1	38.8	44.4
54.3	55.9	64.3	65.2
87.8	88.3	92.6	94.9
115.5	116.9	114.8	121.4
162.5	142.7	126.7	103.3
169.2	177.1	206.7	225.4
263.8	292.2	313.5	321.4
395.2	404.5	411.4	441.2

TABLE I
RECORDED ECONOMIC, DEMOGRAPHIC AND HYDROLOGIC DATA

	A21	GROSS FARM PRODUCT (BILLIONS OF DOLLARS) 1958-1966)	1958 DOLLARS)	1889-1966)	78	14.1	14.7
	11.7	11.5	12.0	11.1	11.5	12.2	13.1
	14.7	14.9	14.8	14.6	15.1	15.5	16.5
	15.7	16.2	15.3	18.1	15.7	17.5	18.8
	16.6	16.4	15.4	16.5	17.6	16.7	16.5
	18.4	17.4	19.7	19.2	18.7	15.6	17.9
	19.5	18.9	20.4	22.2	21.1	20.8	19.4
	19.9	20.9	19.9	20.6	21.6	19.5	20.5
	22.7	23.5	23.9	23.8	24.5	22.5	22.5
					24.1	22.4	21.8
					25.7	24.2	
	35.7	37.2	38.1	40.8	38.7	35.7	39.8
	46.5	48.3	50.7	55.5	57.2	52.6	59.9
	64.4	67.8	67.2	70.2	70.4	66.5	69.1
	80.8	82.5	58.0	65.1	76.4	68.6	71.8
	76.9	65.4	53.3	42.6	45.7	48.1	54.1
	62.7	69.1	88.8	109.8	129.9	128.3	102.5
	92.9	99.6	107.7	109.7	114.1	104.5	100.0
	107.1	107.3	104.3	109.0	110.0	113.5	118.5
						126.6	125.7
							128.2
	181.3	183.9	185.5	187.2	189.0	190.7	192.5
	199.3	201.1	202.4	203.7	205.0	206.1	207.4
	212.5	213.8	218.2	221.0	218.2	225.0	220.4
	223.9	227.4	209.8	217.1	218.6	221.0	225.6
	219.5	217.2	222.0	214.2	213.7	191.7	199.5
	196.0	196.8	190.0	197.6	196.0	194.1	181.1
	162.7	159.0	141.9	134.7	125.3	121.4	123.4
	99.8	96.0	91.7	89.5	84.6	81.3	79.2
	63.6						72.0

TABLE 1
RECORDED ECONOMIC, DEMOGRAPHIC AND HYDROLOGIC DATA

A160	NUMBER OF OPERATING BUSINESSES (THOUSAND, 1870=1970)		101	81	81
	1944	1956			
427	457	500	603	639	661
747	782	822	920	979	1047
1111	1143	1193	1209	1152	1051
1174	1219	1253	1281	1320	1148
1515	1525	1564	1617	1655	1486
1821	1927	1983	1996	2047	1486
2183	2125	2077	1961	1974	1708
2156	2171	2152	2023	1855	1733
2687	2698	2637	2667	2632	2199
2708	2641	2589	2544	2524	2213
2442					2116
					2192
					2550
					2679
					2708
					2444
A163	OUTPUT PER UNIT OF LABOR INPUT INDEX (1958=100, 1889=1969)				
26.8	28.0	28.5	30.0	28.9	30.1
32.6	33.8	34.9	33.1	34.9	32.1
37.2	36.2	36.9	37.3	36.3	35.1
43.8	42.6	46.4	45.4	47.2	36.6
53.5	52.3	54.9	53.2	52.1	37.6
63.5	64.9	66.3	65.7	65.8	40.1
76.8	82.3	82.9	83.8	87.1	41.3
103.8	104.5	104.5	107.7	112.4	51.3
132.0					51.1
					59.3
					61.5
					73.5
					73.5
					100.0
					131.1
A165	OUTPUT PER UNIT OF CAPITAL INPUT INDEX (1958=100, 1889=1969)				
58.0	60.3	59.8	61.7	55.7	53.6
60.4	59.7	64.5	62.5	60.4	57.6
65.3	63.9	64.3	66.0	66.6	56.0
67.3	66.2	62.9	66.4	74.3	59.1
77.5	68.1	63.1	56.2	56.8	66.7
83.9	89.3	90.2	102.9	73.9	74.5
102.4	109.1	109.5	106.7	74.7	75.6
104.5	104.1	103.5	108.0	62.1	76.4
112.0				69.8	81.6
				15.6	105.6
				15.7	105.6
				109.4	109.4
				108.4	108.4
				103.2	103.2
				111.9	111.9
				114.5	114.5
				115.3	115.3
				112.7	112.7

TABLE 1
RECORDED ECONOMIC, DEMOGRAPHIC AND HYDROLOGIC DATA

B1	UNEMPLOYMENT RATE (1890-1971)	82
4.00	5.40	12.4
5.00	5.40	6.50
5.00	4.00	6.00
5.90	6.70	5.10
5.20	11.7	1.40
6.70	6.70	1.40
6.70	4.60	1.20
11.7	4.30	4.20
15.9	2.40	3.20
15.9	23.6	1.80
14.6	4.70	21.7
14.6	9.90	24.1
5.00	3.30	21.9
5.50	6.70	16.9
4.90	5.90	14.3
		19.0
		17.2
		3.80
		5.90
		5.50
		6.80
		3.60
		3.50
B50	APPLICATIONS FOR PATENTS (1860-1970)	111
7653	4643	15269
19171	19472	21276
22395	25556	20420
40930	40443	19271
41898	46334	20260
64448	68904	20308
86575	93063	35461
93752	83967	35684
69484	59609	40464
74108	64788	41337
84246	87921	61273
109052		65642
		80337
		59581
		92364
		94272
		75006
		71306
		75847
		74660
		82552
		83587
		98402
		103993

TABLE 1
RECORDED ECONOMIC, DEMOGRAPHIC AND HYDROLOGIC DATA

		111	
		111	
B68	WHOLE SALE PRICE INDEX (1967=100, 1860=1970)	56.5	53.6
34.8	32.8	38.1	48.5
47.5	45.4	48.4	46.5
35.8	35.5	37.0	34.3
28.9	28.8	26.9	27.5
36.4	33.5	36.0	35.6
79.6	58.3	49.9	51.9
44.6	37.6	33.6	34.0
48.5	45.1	45.1	41.9
81.8	91.1	88.6	87.4
94.9	94.5	94.8	94.5
110.4			
B69	CONSUMER PRICE INDEX (1967=100, 1860=1970)	44.0	43.0
27.0	27.0	48.0	47.0
38.0	37.0	37.0	35.0
30.0	30.0	29.0	28.0
28.0	28.0	28.0	26.0
25.0	25.0	26.0	27.0
29.0	29.0	30.0	29.7
60.0	53.6	50.2	51.1
50.0	45.6	49.9	48.8
42.0	42.0	44.1	48.8
72.1	72.1	77.8	79.5
88.7	88.7	89.6	90.6
116.3			

TABLE 1
RECORDED ECONOMIC, DEMOGRAPHIC AND HYDROLOGIC DATA

877	MANHATTAN ISLAND REAL ESTATE MORTGAGE RATES (1879-1970)	92	5.20
5.92	5.78	5.65	5.43
5.13	5.45	5.38	5.20
5.05	5.17	5.11	5.09
5.35	5.35	5.47	5.46
5.65	5.75	5.97	5.95
5.92	5.95	5.75	5.77
5.05	5.03	4.90	4.98
4.93	4.95	4.93	5.03
5.71	5.85	5.87	5.90
7.06	7.73		
879	STOCK EXCHANG CALL LOAN RATES (1860-1970)	111	10.29
5.99	5.76	5.23	5.23
5.72	5.56	5.38	5.43
4.86	5.76	4.78	5.71
5.84	3.42	3.98	4.57
2.94	4.00	5.15	3.71
2.98	2.57	3.52	3.22
7.74	5.97	4.29	4.86
2.94	1.74	2.05	1.16
1.00	1.00	1.00	1.00
1.63	2.17	2.48	3.06
4.99	4.58	4.50	4.50
7.95			7.95
880	COMMERCIAL PAPER RATES (1890-1970)	81	5.50
6.91	6.48	5.40	7.64
5.71	5.40	5.81	6.16
5.72	4.75	5.41	6.20
7.50	6.62	4.52	5.07
3.59	2.64	2.73	1.73
8.56	8.54	8.66	8.69
1.45	2.16	2.33	1.58
3.85	2.97	3.26	3.55
7.72			3.97

TABLE 1
RECORDED ECONOMIC, DEMOGRAPHIC AND HYDROLOGIC DATA

ENR CONSTRUCTION COSTS INDEX (1913=100, 1903=1973)						
	94	87	91	95	101	97
100	89	93	130	181	189	198
214	215	207	208	206	207	207
170	196	196	206	235	236	242
290	299	306	346	413	461	477
600	628	660	692	724	759	797
901	936	971	1019	1070	1155	1269
1895						

C71 POPULATION-UNITED STATES (THOUSANDS, 1900=1970)						
	76094	77585	79160	80632	82165	83826
92407	93868	95351	97227	99118	100549	101966
106466	108541	110055	111950	114113	115832	117399
123077	124948	124849	125579	126374	127250	128053
131954	133121	133920	134245	132885	132481	140054
151235	153310	155687	158241	161164	164308	167305
179975	182973	185738	188438	191085	193460	195501
203805						

C72 POPULATION-NEW ENGLAND DIVISION (THOUSANDS, 1900=1970)						
	5579	5624	5719	5809	5906	6009
6565	6629	6729	6883	7034	7137	7198
7454	7574	7668	7756	7852	7946	8018
8175	8193	8220	8254	8296	8361	8391
8450	8556	8572	8404	8335	8321	8839
9315	9288	9358	9627	9832	9871	9929
10532	10666	10808	10986	11185	11329	11431
11873						

TABLE 1
RECORDED ECONOMIC, DEMOGRAPHIC AND HYDROLOGIC DATA

C73 POPULATION-MIDDLE ATLANTIC DIVISION (THOUSANDS, 1900-1970)							71
15480	15777	16096	16439	16810	17205	17638	18066
19393	19709	20017	20341	20650	20912	21229	21547
22220	22613	22951	23372	23909	24328	24609	24810
26364	26675	26885	27017	27137	27234	27332	27389
27526	27432	27003	26477	26000	25746	27755	28796
30243	30357	30821	31418	31991	32407	32699	33065
34270	34717	35033	35416	35768	36122	36358	36544
37271							36941

C74 POPULATION-EAST NORTH CENTRAL DIVISION (THOUSANDS, 1900-1970)							71
16002	16230	16499	16727	16969	17216	17435	17654
18344	18614	18918	19321	19748	20116	20452	20771
21811	22349	22680	23159	23667	24053	24390	24792
25332	25426	25535	25632	25694	25824	25960	26096
26725	27049	27135	26498	26524	26381	28411	29151
30530	30884	31498	32167	33063	33780	34404	34967
36291	36615	36926	37357	37869	38406	38953	39347
40368							39945

C75 POPULATION-WEST NORTH CENTRAL DIVISION (THOUSANDS, 1900-1970)							71
10357	10492	10627	10719	10861	10987	11089	11214
11675	11771	11868	12006	12166	12316	12460	12474
12562	12656	12734	12812	12904	12986	13071	13145
13335	13446	13518	13567	13593	13630	13601	13544
13498	13285	13101	12786	12481	12427	13165	13447
14102	14169	14175	14269	14453	14731	14905	14978
15424	15570	15657	15716	15786	15819	15887	15942
16367							16047

TABLE 1
RECORDED ECONOMIC, DEMOGRAPHIC AND HYDROLOGIC DATA

C76	POPULATION=SOUTH ATLANTIC DIVISION	(THOUSANDS, 1900-1970)	71
10456	10625	10787	10937
12259	12474	12669	12954
14161	14340	14484	14649
15816	15942	16066	16205
17935	18439	18935	19460
21258	21796	22134	22384
26091	26668	27183	27742
38772			

C77	POPULATION=EAST SOUTH CENTRAL DIVISION	(THOUSANDS, 1900-1970)	71
7554	7685	7775	7858
8442	8516	8599	8737
8989	9038	9144	9265
9895	9952	10044	10137
10814	10883	10888	10847
11486	11534	11492	11380
12073	12198	12316	12416
12823			

C78	POPULATION=WEST SOUTH CENTRAL DIVISION	(THOUSANDS, 1900-1970)	71
6553	6801	7037	7260
8643	9035	9169	9375
10347	10577	10738	10923
12209	12282	12346	12438
13875	13388	13448	13626
14618	14991	15214	15164
17010	17292	17678	17650
19397			

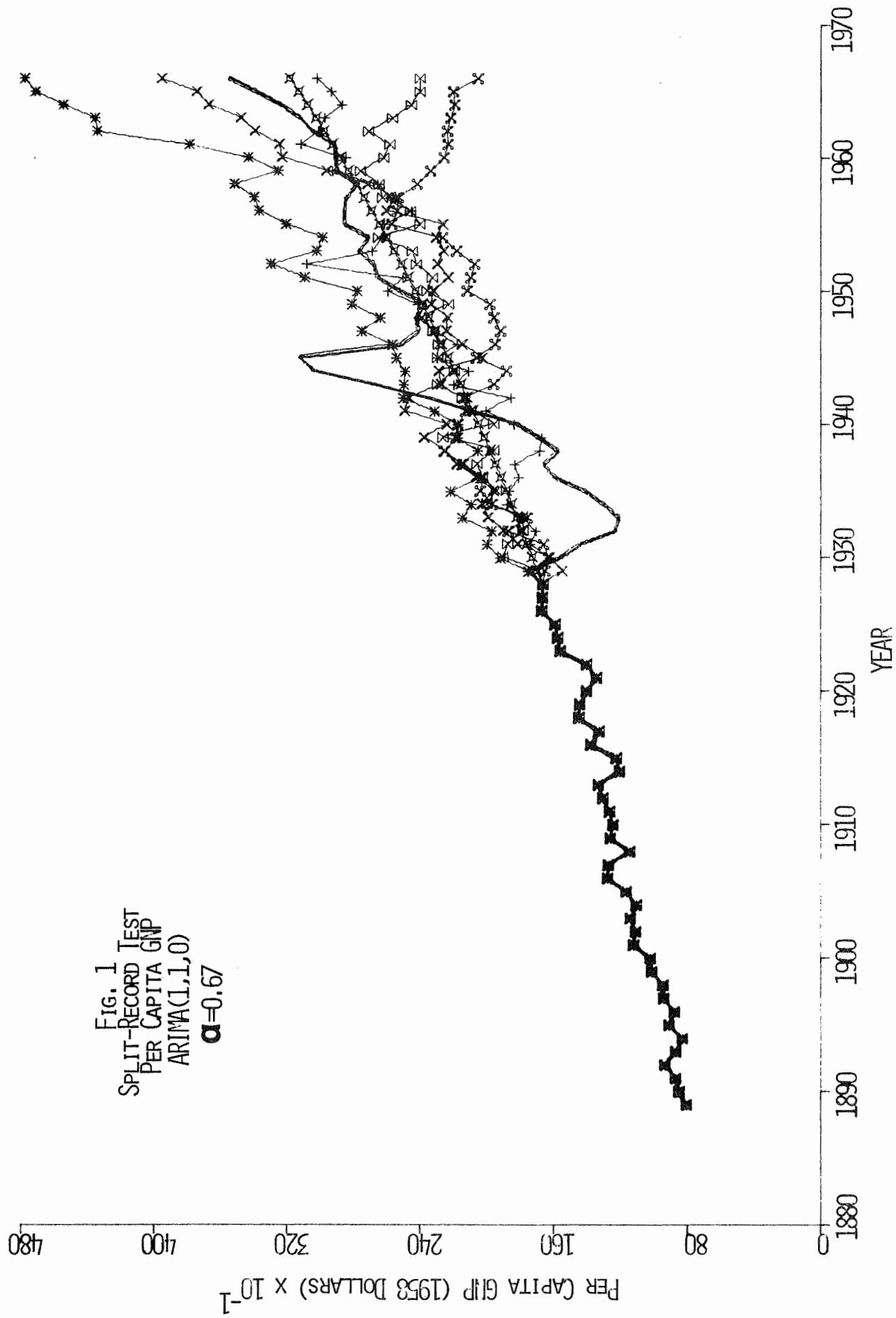
TABLE I
RECORDED ECONOMIC, DEMOGRAPHIC AND HYDROLOGIC DATA

C79	POPULATION=MOUNTAIN DIVISION (THOUSANDS, 1900-1970)	71	2435	2541
1684	1784	1892	2049	2236
2654	2725	2793	2868	3112
3344	3403	3435	3478	3567
3714	3751	3775	3797	3835
4155	4082	4098	4393	4272
5102	5179	5330	5515	5670
6916	7183	7381	7541	7662
8345			7740	7802
			7878	8008
			8171	
C80	POPULATION=PACIFIC DIVISION (THOUSANDS, 1900-1970)	71	3799	4016
2428	2567	2728	2900	3078
4231	4393	4558	4742	4901
5715	5998	6220	6536	6864
8235	8371	8450	8533	8655
9776	10096	10743	11754	12271
14596	15113	15665	16318	16889
21368	22063	22761	23414	23963
26589			24464	24865
			25329	25688
			26162	
MEAN ANNUAL FLOW=COLUMBIA R.(CFS, 1879-1970)	92	201000	260000	201000
242000	264000	252000	232000	212000
134000	196000	165000	198000	219000
234000	224000	219000	217000	211000
190000	213000	188000	183000	213000
172000	157000	230000	183000	179000
133000	131000	122000	186000	198000
149700	148500	130000	178600	207300
180100	217100	226400	198200	179300
211600	195300	189900	169300	174100
205000	166500		184200	224500
			162000	191300
			165700	

RECORDED ECONOMIC, DEMOGRAPHIC AND HYDROLOGIC DATA

MEAN ANNUAL FLOW-ST.	LAWRENCE R. (CFS, 1861-1965)
275000	273000
284000	266000
293000	260000
302000	257000
311000	261000
320000	251000
329000	242000
338000	242000
347000	244000
356000	241000
365000	258000
374000	229000
383000	234000
392000	218000
401000	213000
410000	214000
419000	221000
428000	217000
437000	229000
446000	186000
455000	206000
464000	168000
473000	214000
482000	183000
491000	214000
500000	192000
509000	208000
518000	214000
527000	206000
536000	214000
545000	221000
554000	217000
563000	246000
572000	264000
581000	248000
590000	210000
600000	276000
610000	213000
620000	229900
630000	210300
640000	205400
650000	206000
660000	191600
670000	223000
680000	236000
690000	241000
700000	216000
710000	218000
720000	255000
730000	263000
740000	226000
750000	231000
760000	241000
770000	228000
780000	246000
790000	259000
800000	252000
810000	242000
820000	272000
830000	249000
840000	279000
850000	115

FIG. 1 TEST
PER CAPITA GNP
ARIMA(1,1,0)
 $\alpha = 0.67$



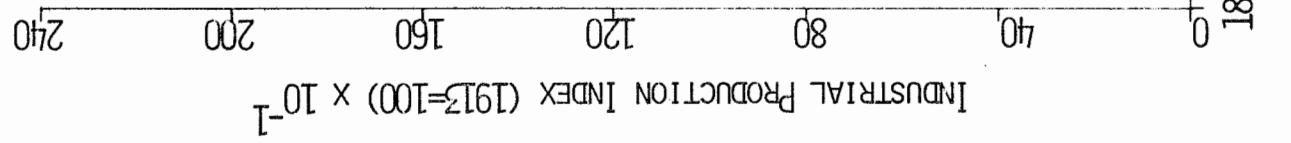


FIG. 2
SPLIT-RECORD TEST
INDUSTRIAL PRODUCTION INDEX
ARIMA(0,1,1)
 $\alpha = 0.94$

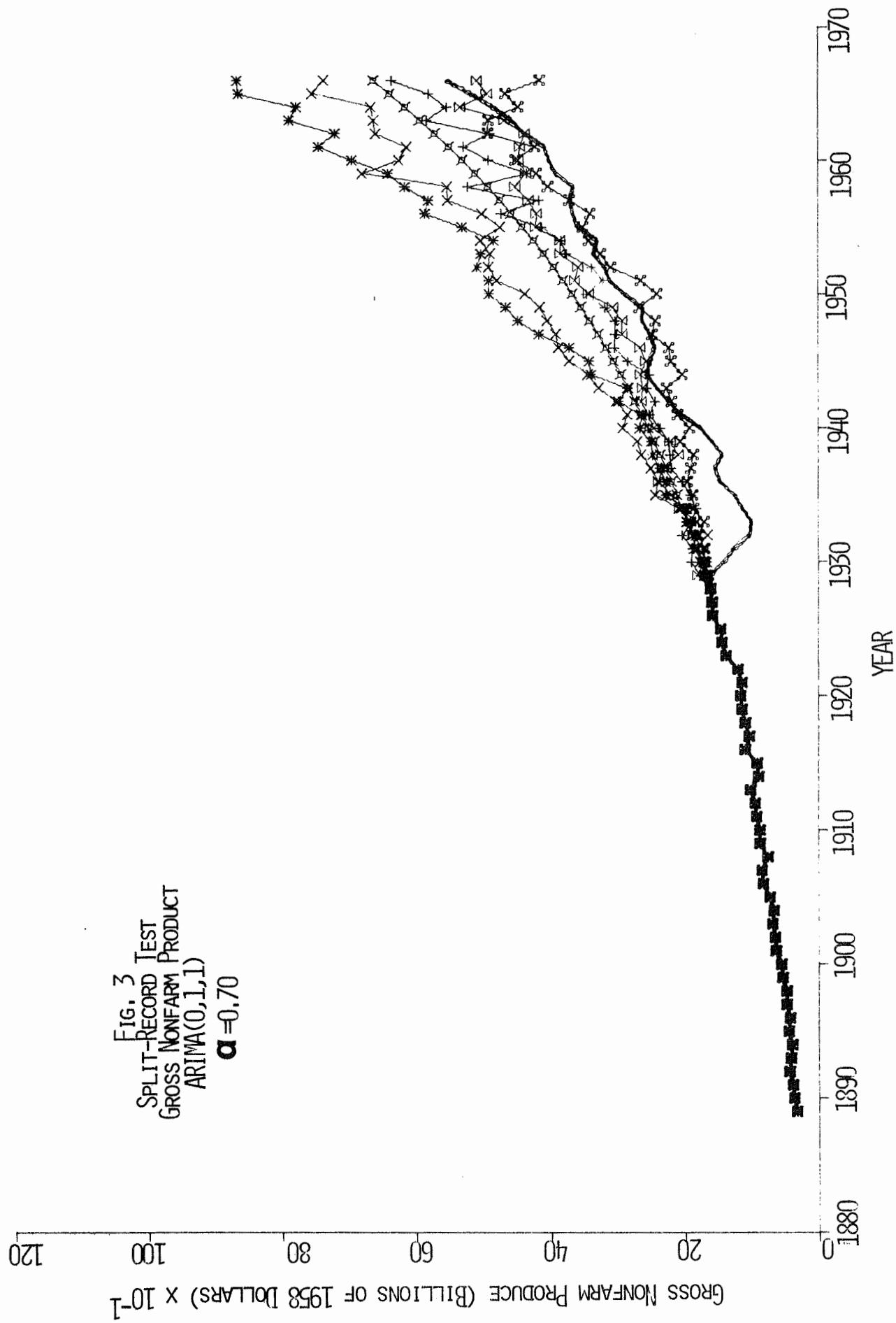




FIG. 5
SPLIT-RECORD TEST
MAN-HOURS IN MANUFACTURING INDEX
ARIMA(0,1,1)
 $\alpha=0.66$

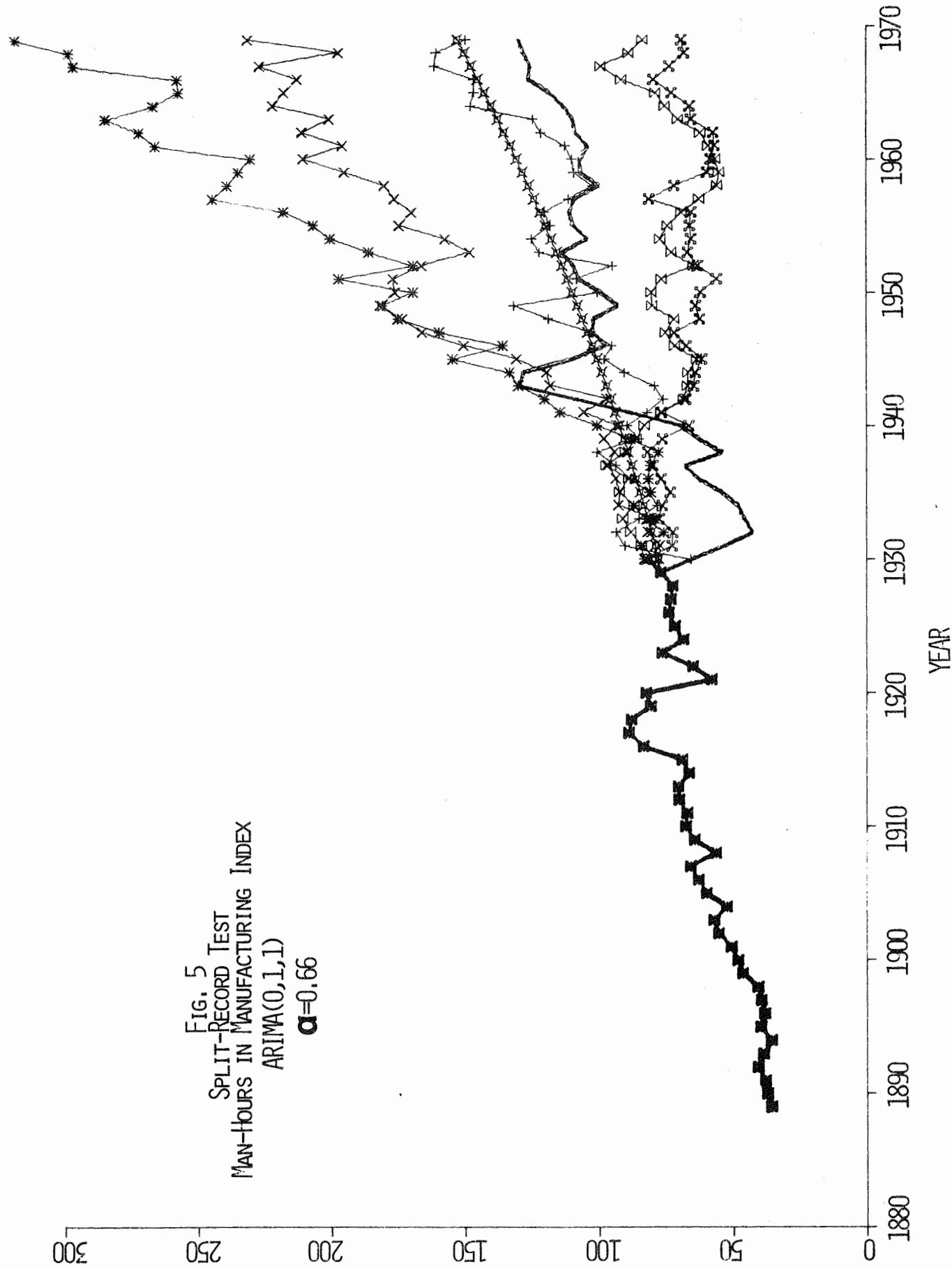
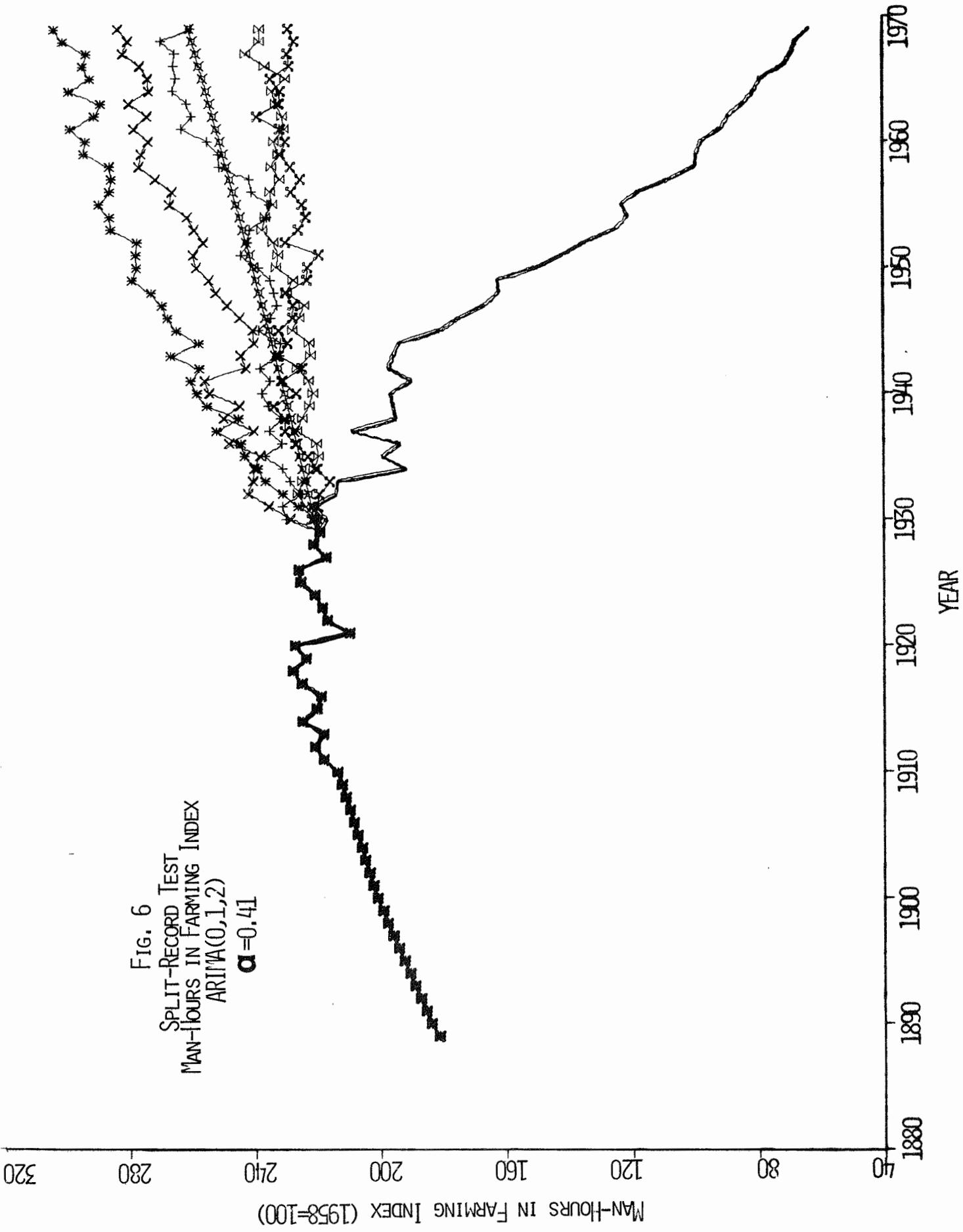


FIG. 6
SPLIT-RECORD TEST
MAN-HOURS IN FARMING INDEX
ARIMA(0,1,2)
 $\alpha = 0.4$



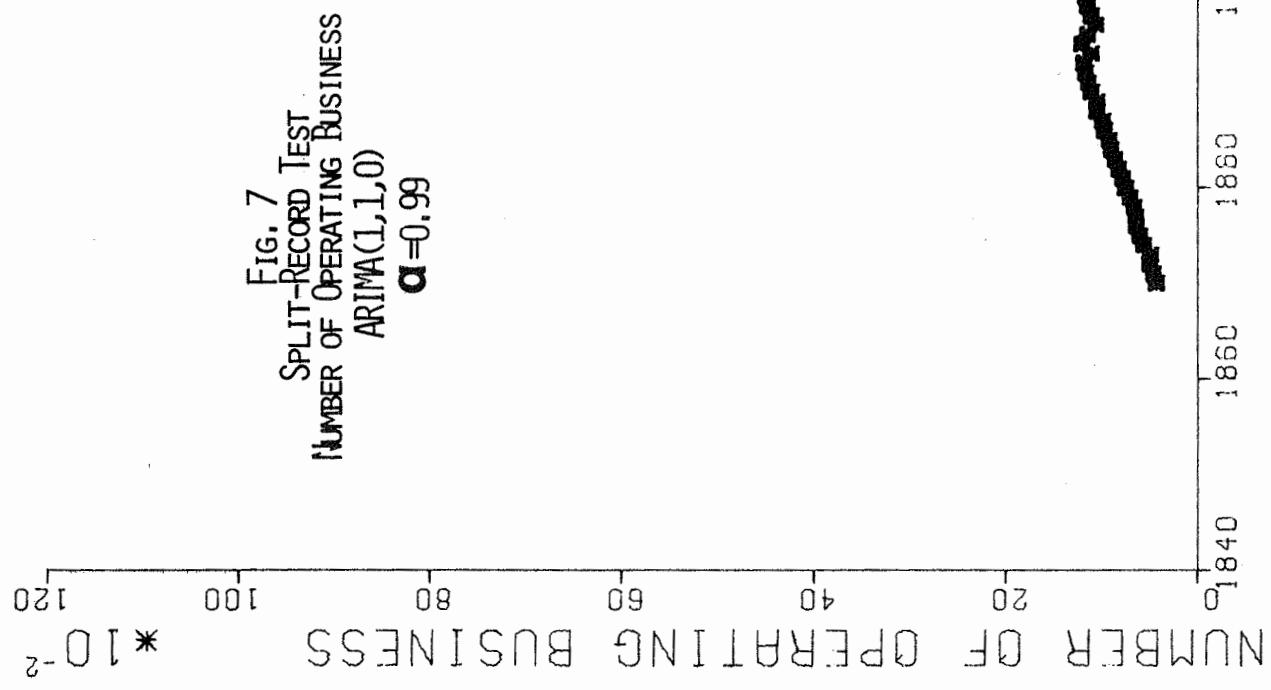
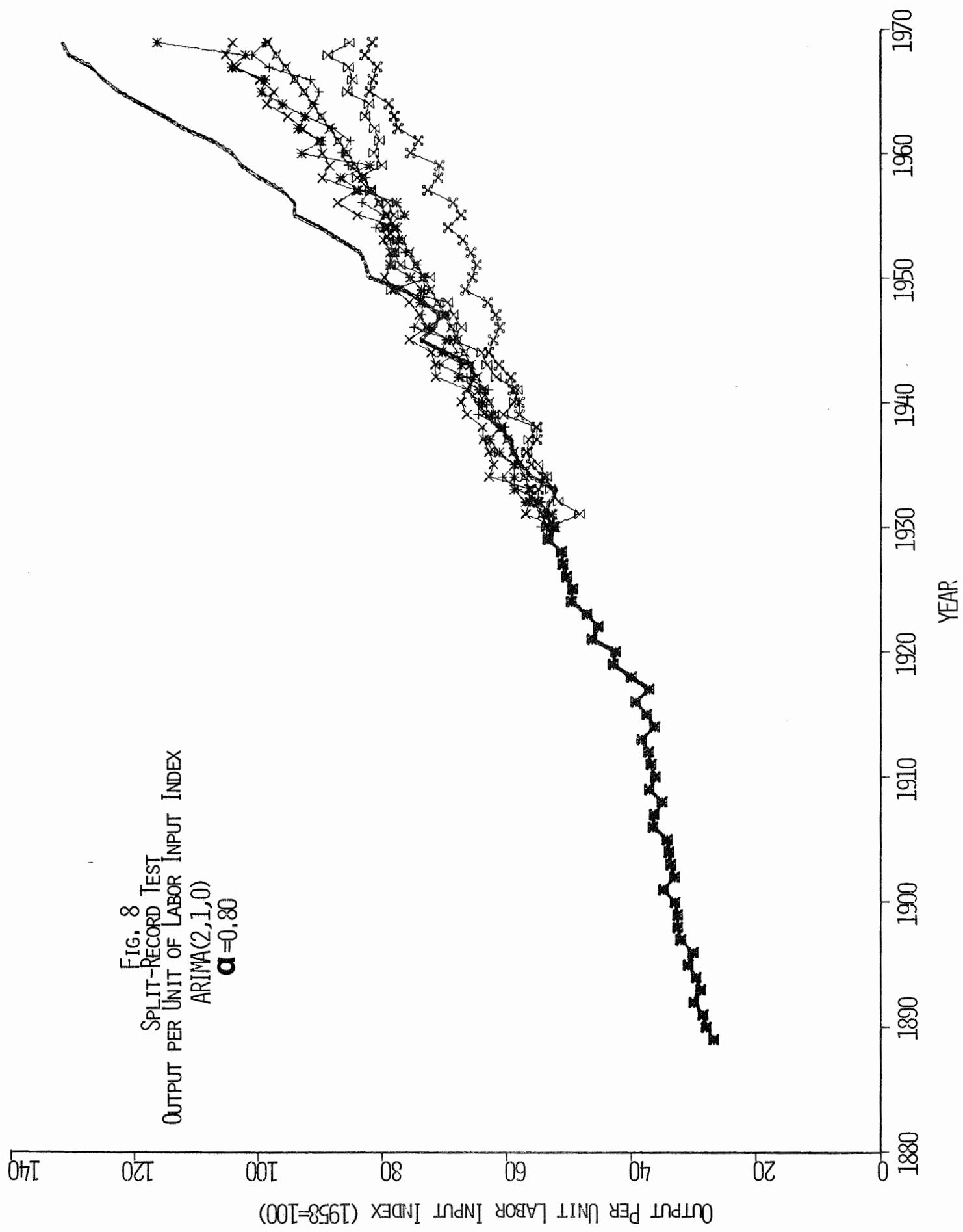
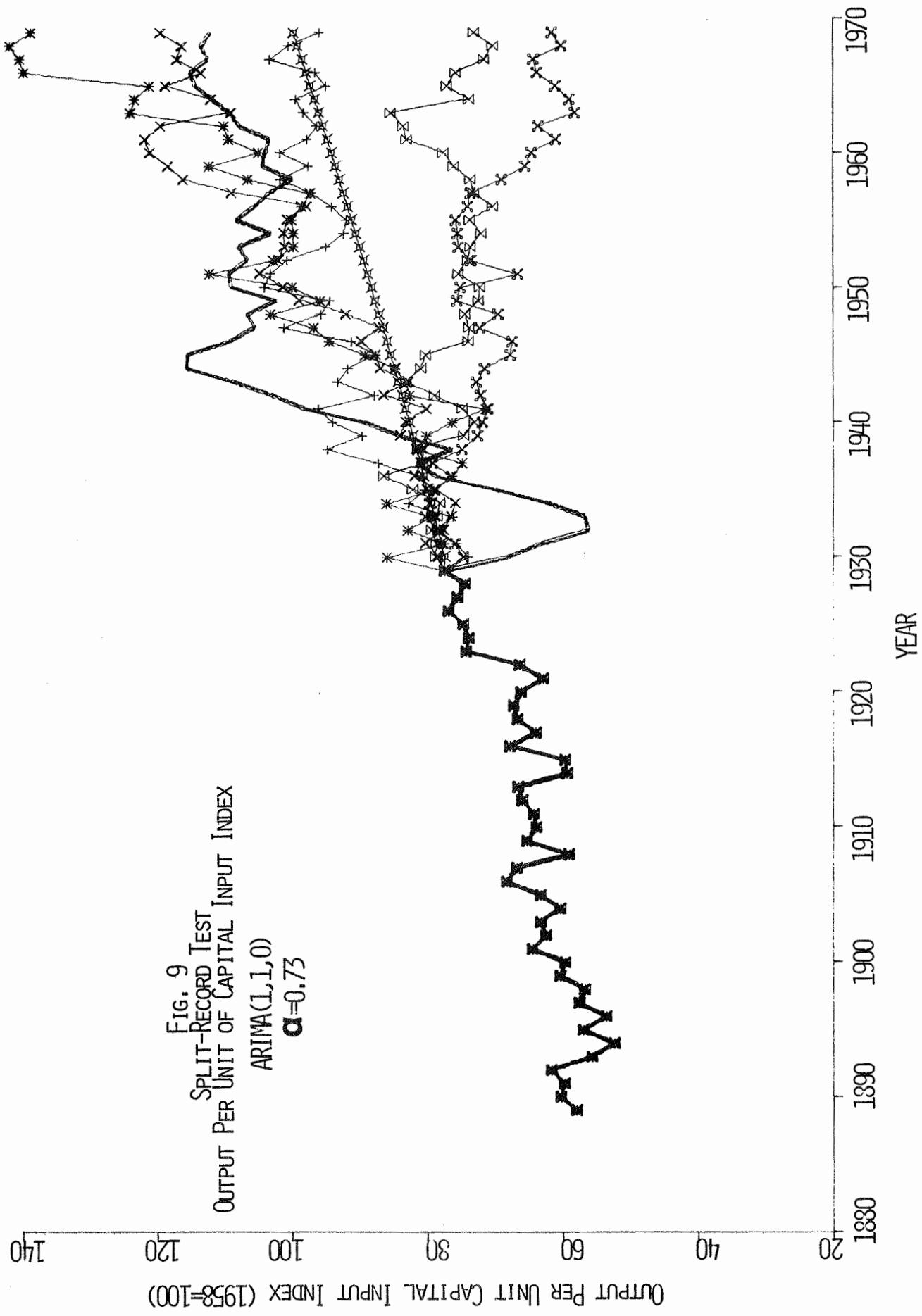
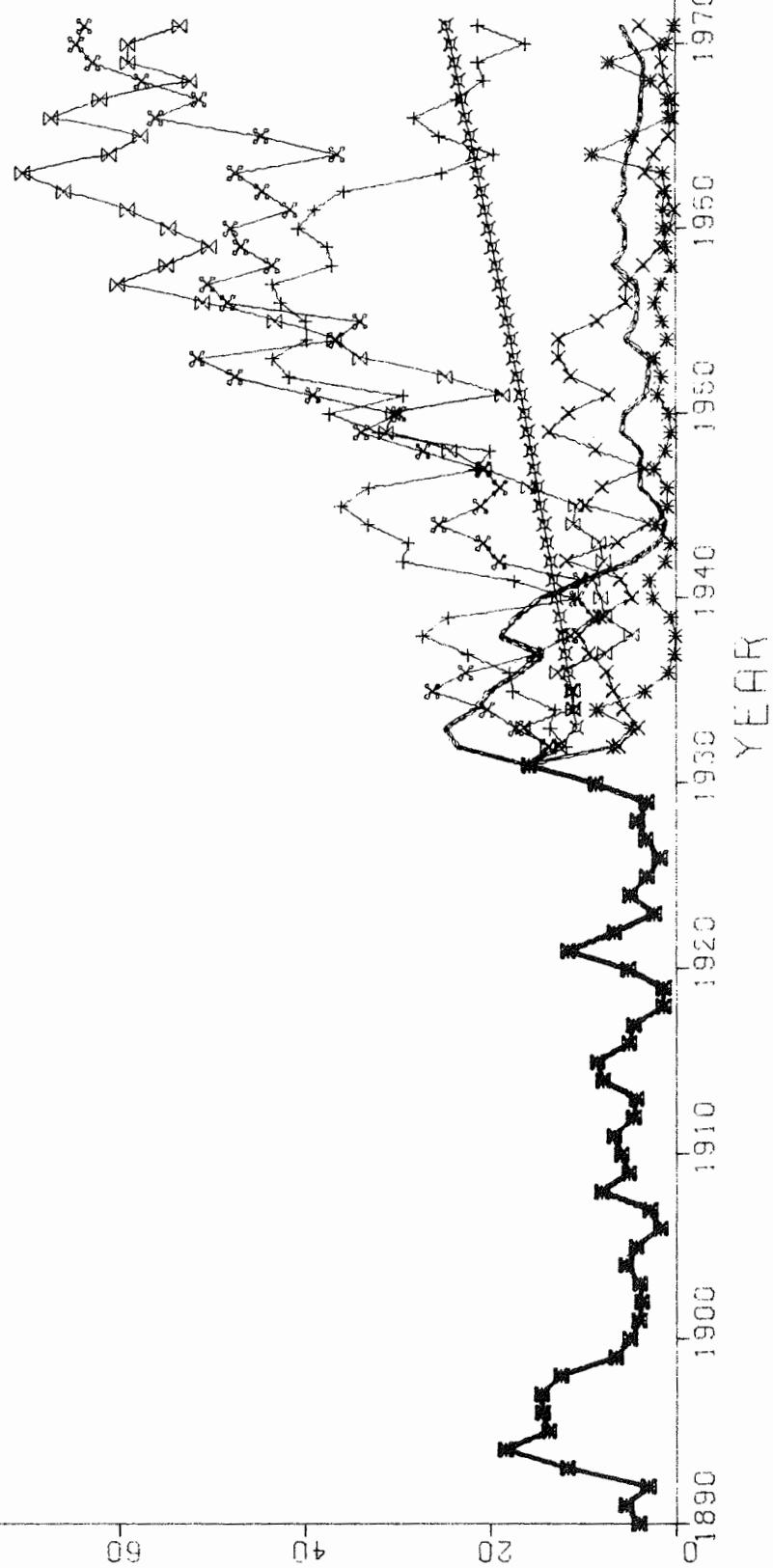
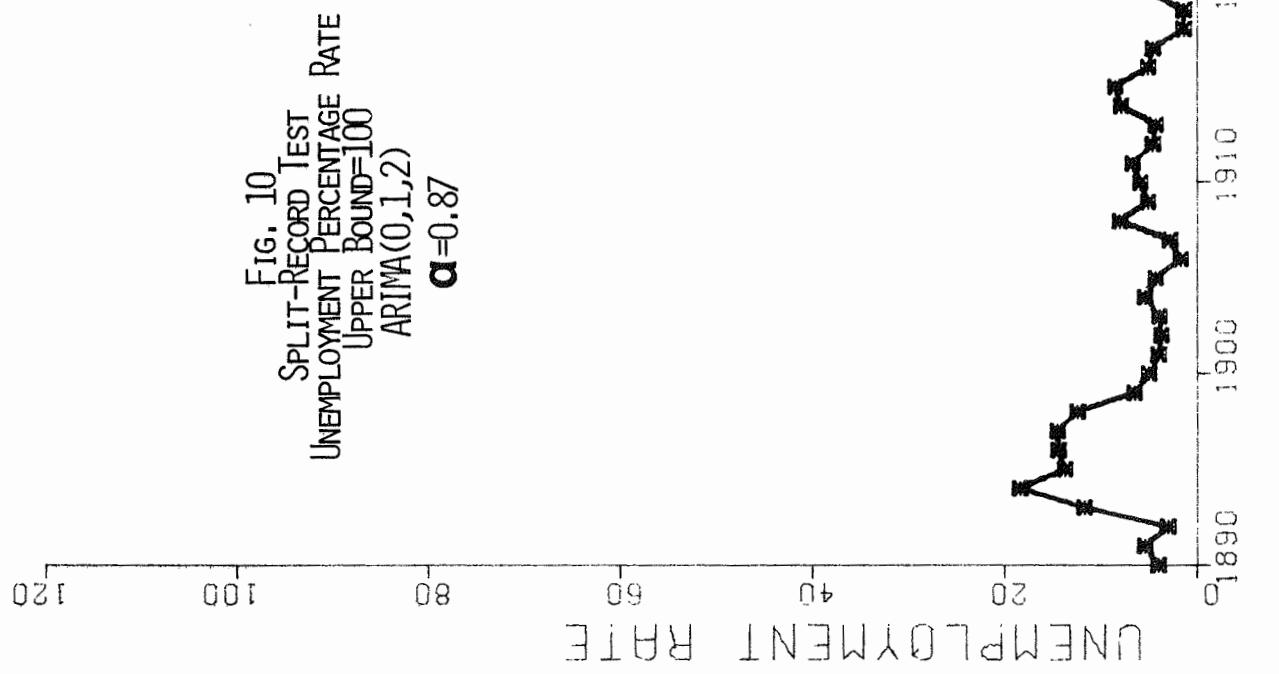
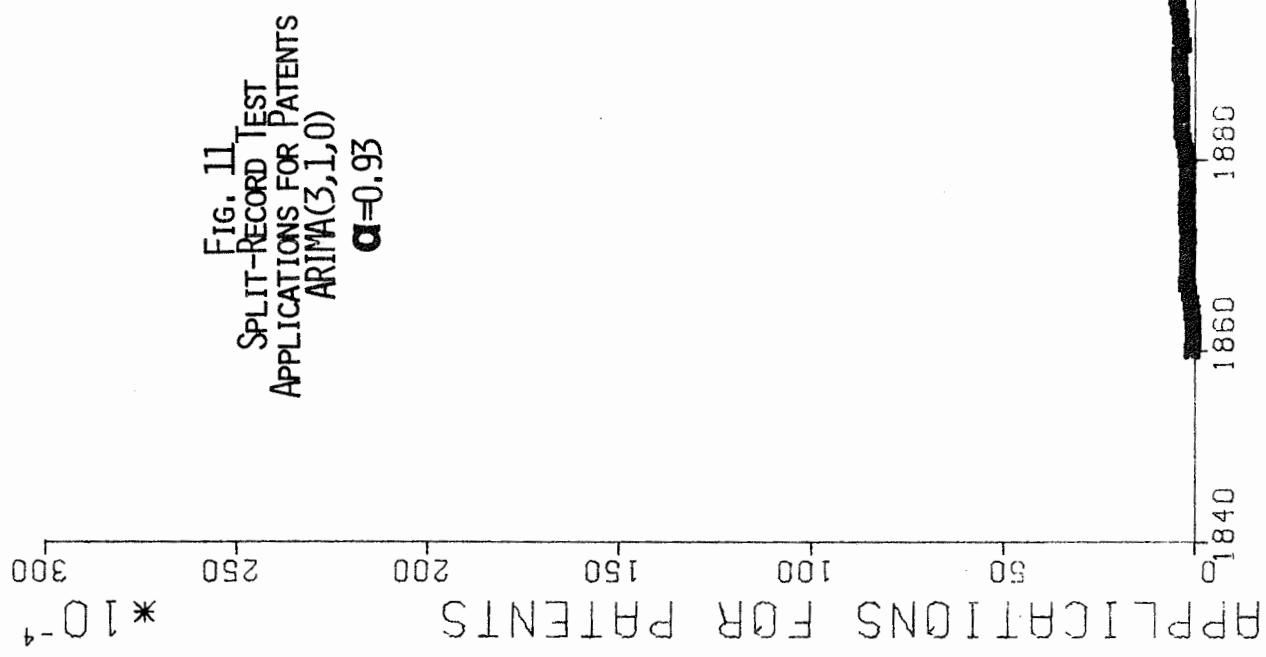


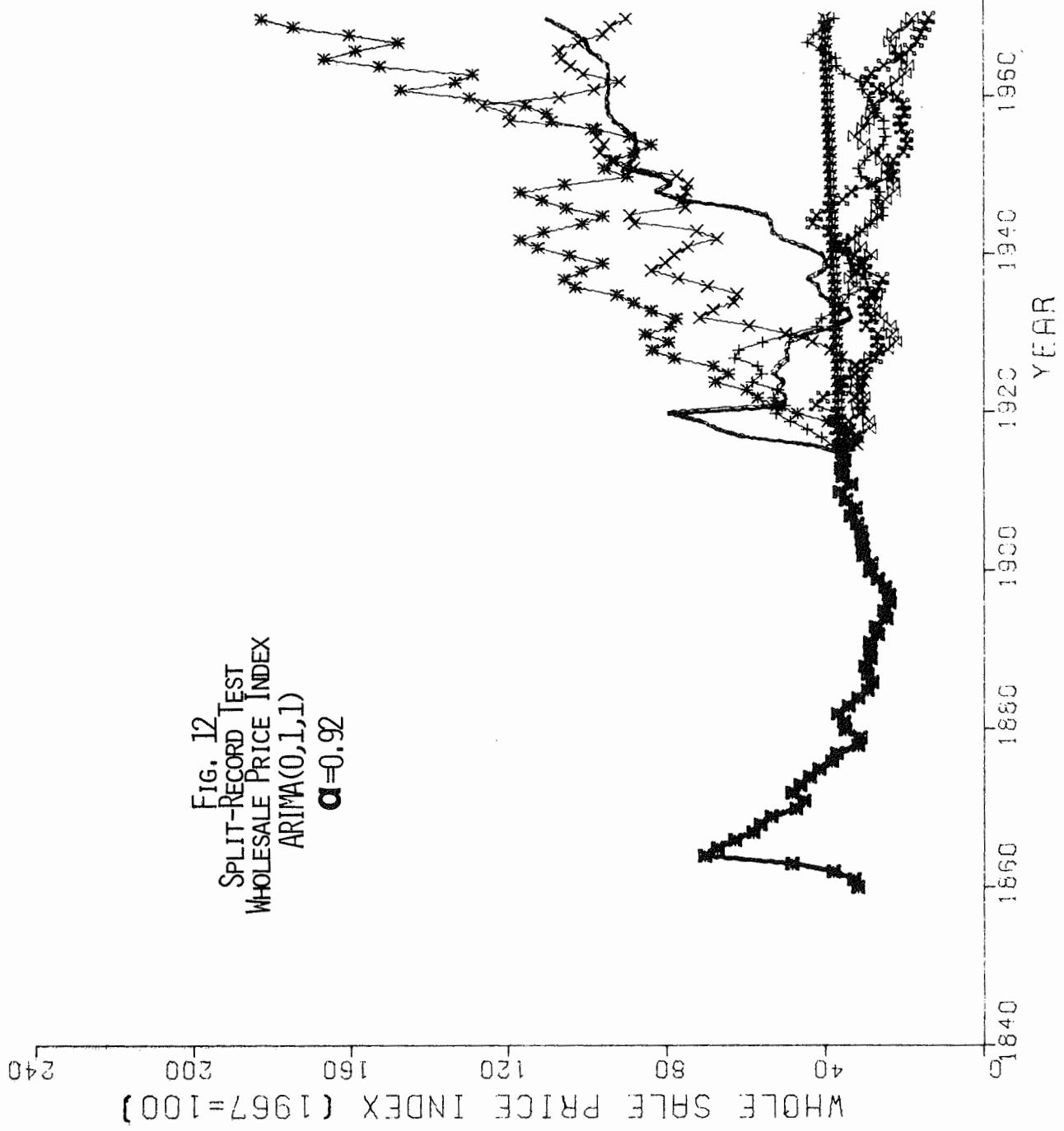
FIG. 8
SPLIT-RECORD TEST
OUTPUT PER UNIT OF LABOR INPUT INDEX
ARIMA(2,1,0)
 $\alpha = 0.80$

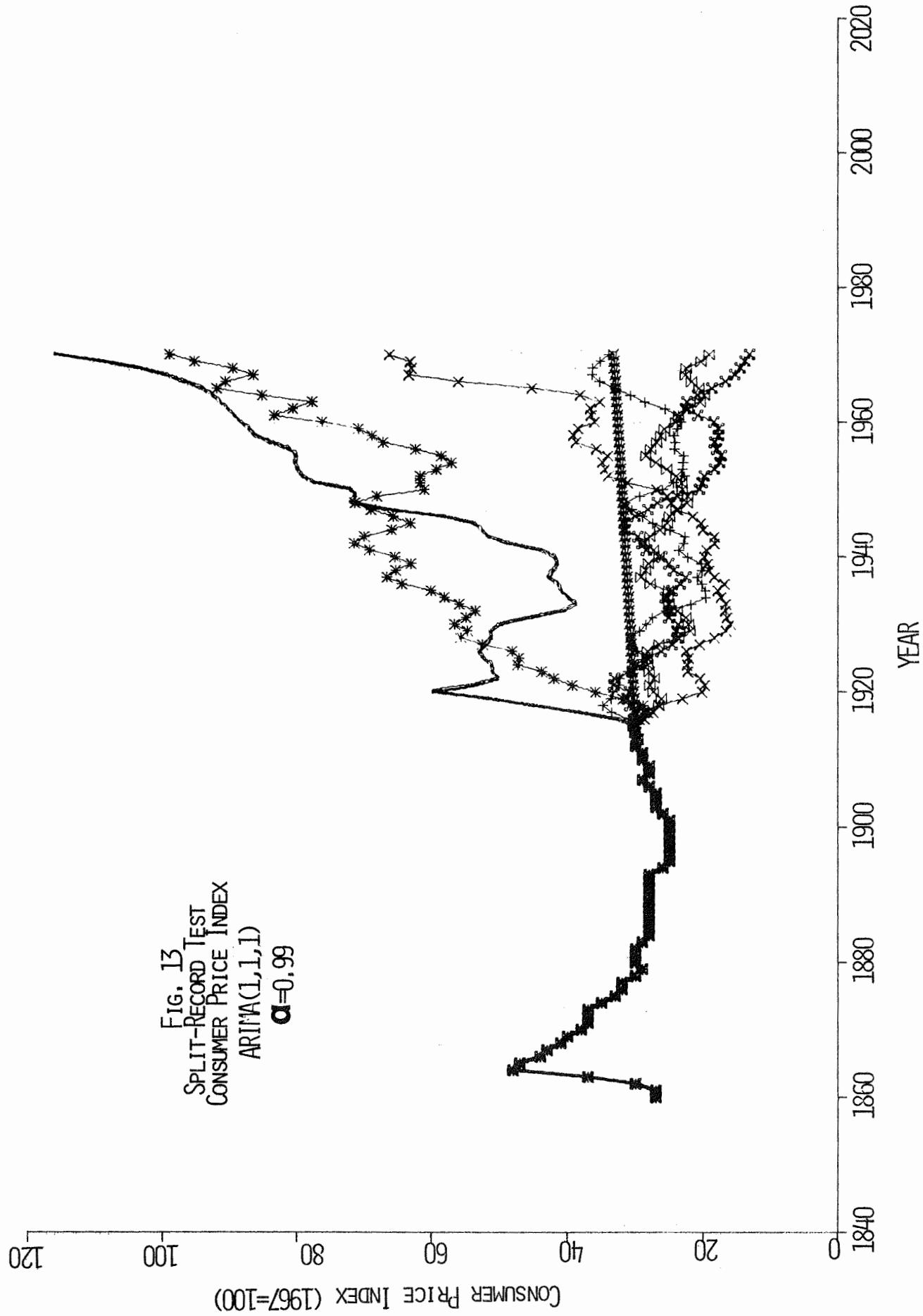












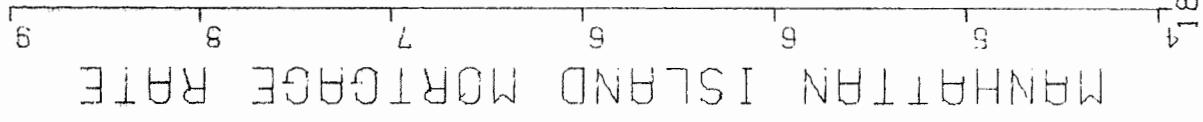


FIG. 14
SPLIT-RECORD TEST
MANHATTAN ISLAND MORTGAGE RATES
ARIMA(0,1,3)
 $\alpha = 0.46$

Fig. 15
SPLIT-RECORD TEST
Stock Exchange CALL LOAN RATES
ARIMA(2,1,0)
 $\alpha = 0.85$

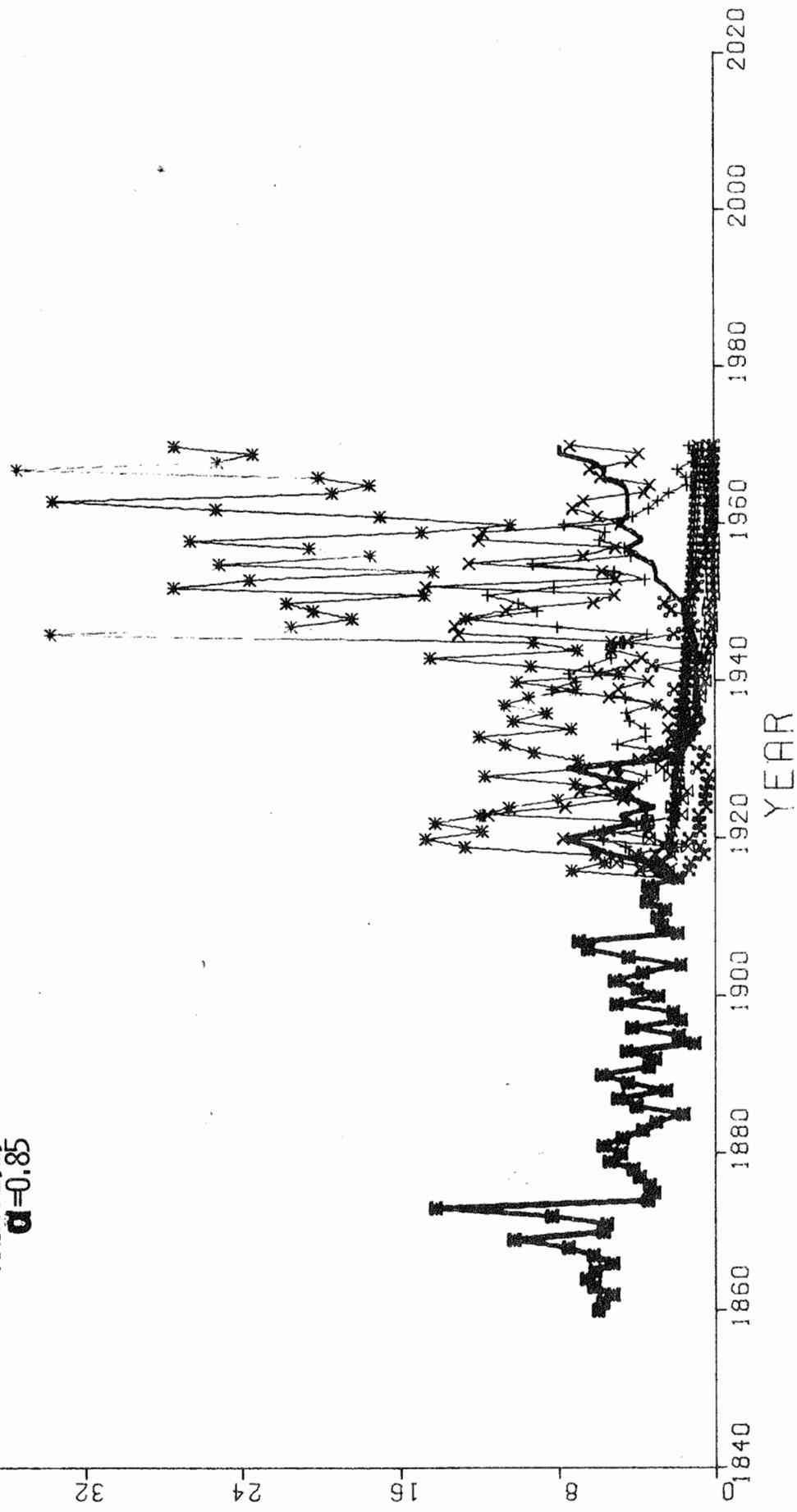


FIG. 16
SPLIT-RECORD TEST
COMMERCIAL PAPER RATES
 $\text{ARIMA}(1,1,0)$
 $\alpha = 0.55$

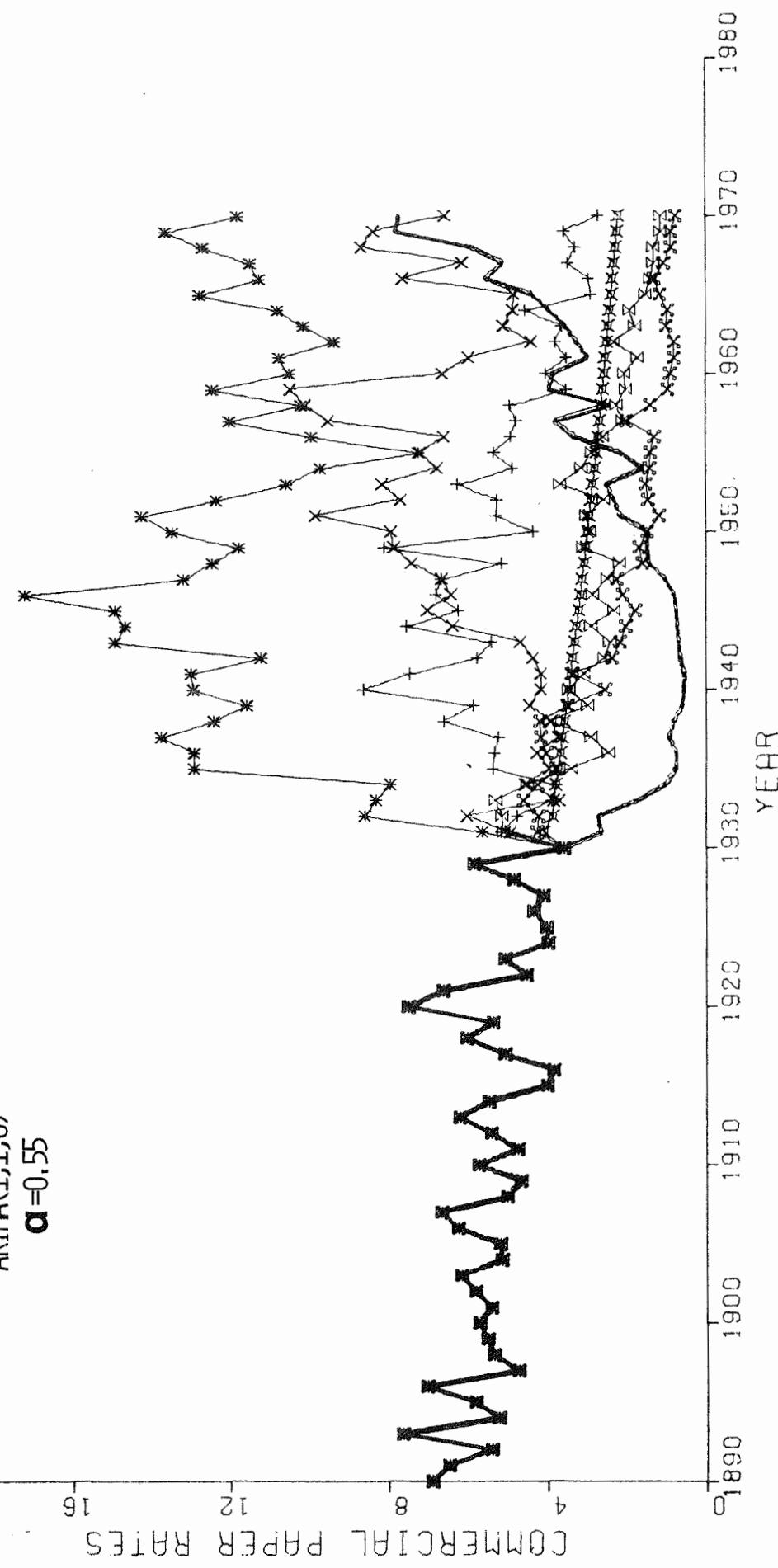
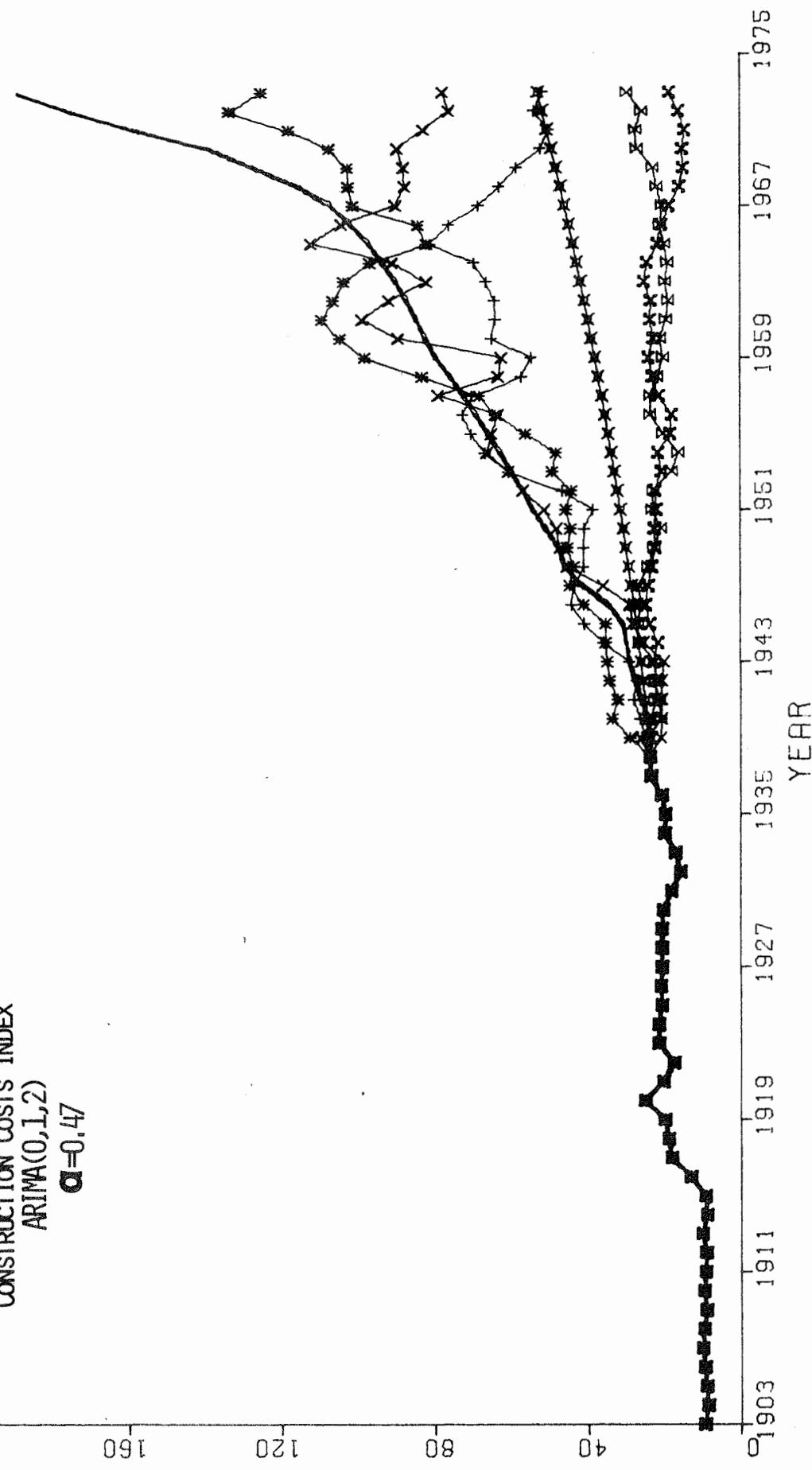
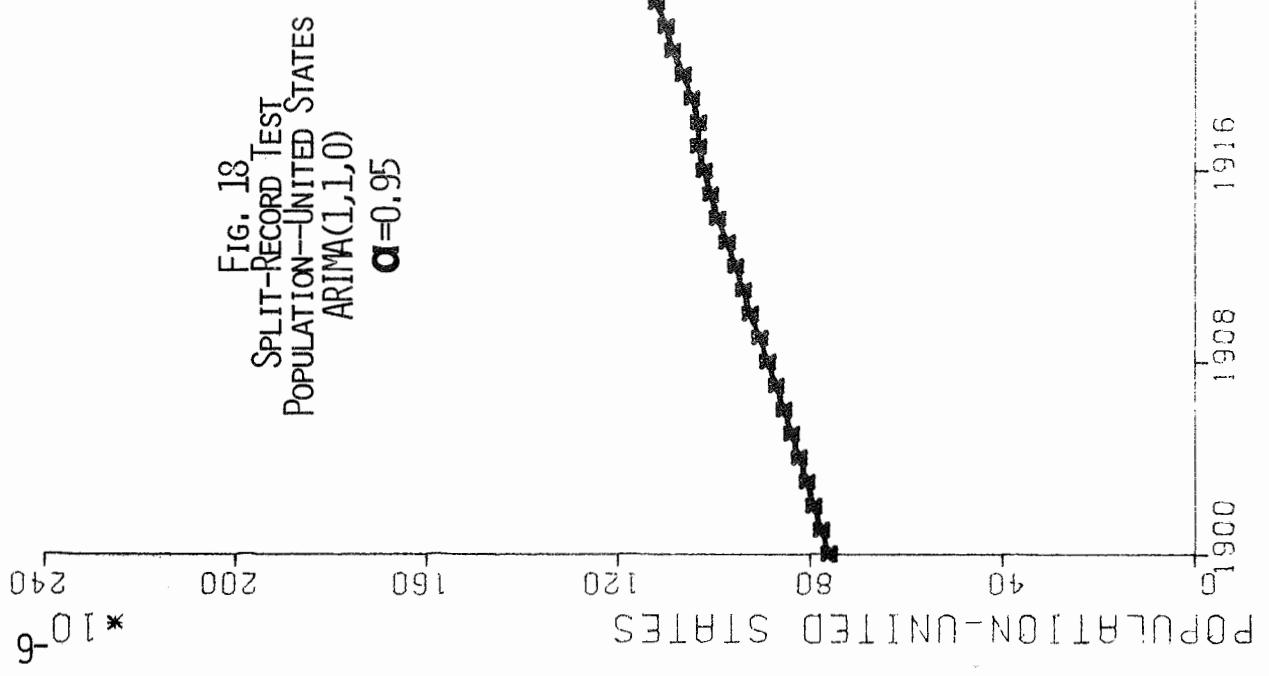


Fig. 17
SPLIT-RECORD TEST
CONSTRUCTION COSTS INDEX
ARIMA(0,1,2)
 $\alpha=0.47$





1900 1908 1916 1924 1932 1940 1948 1956 1964 1972
 YEAR

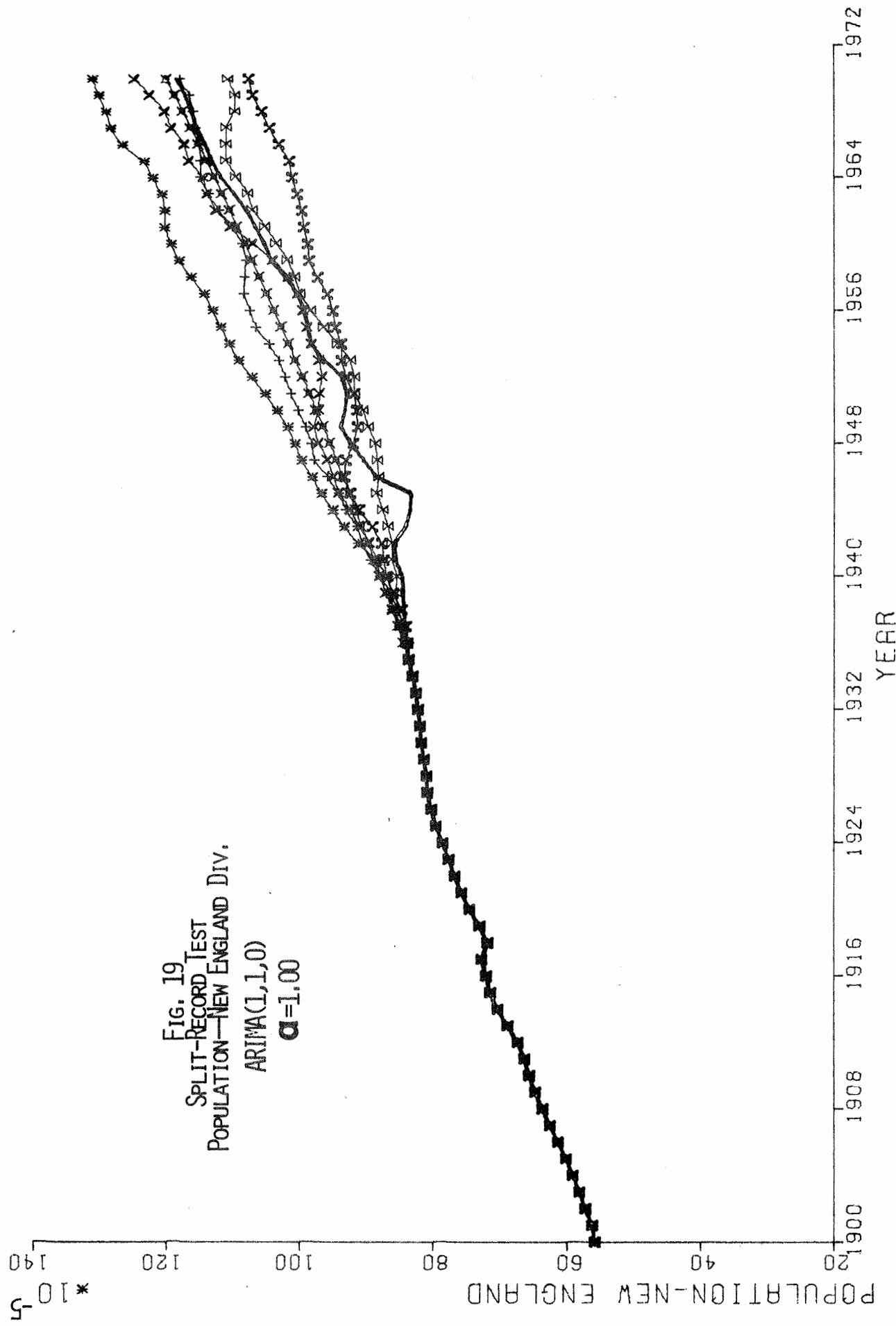
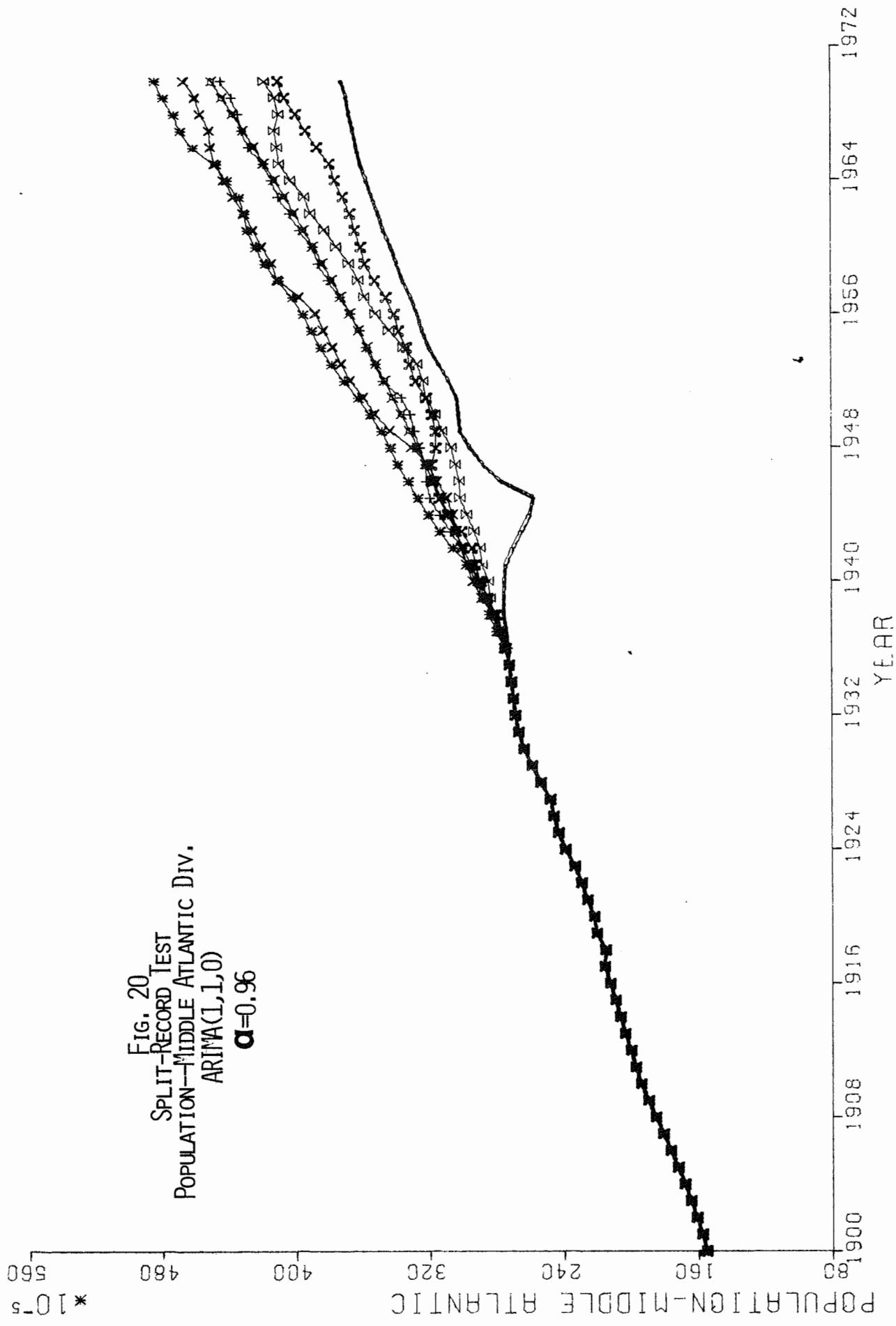
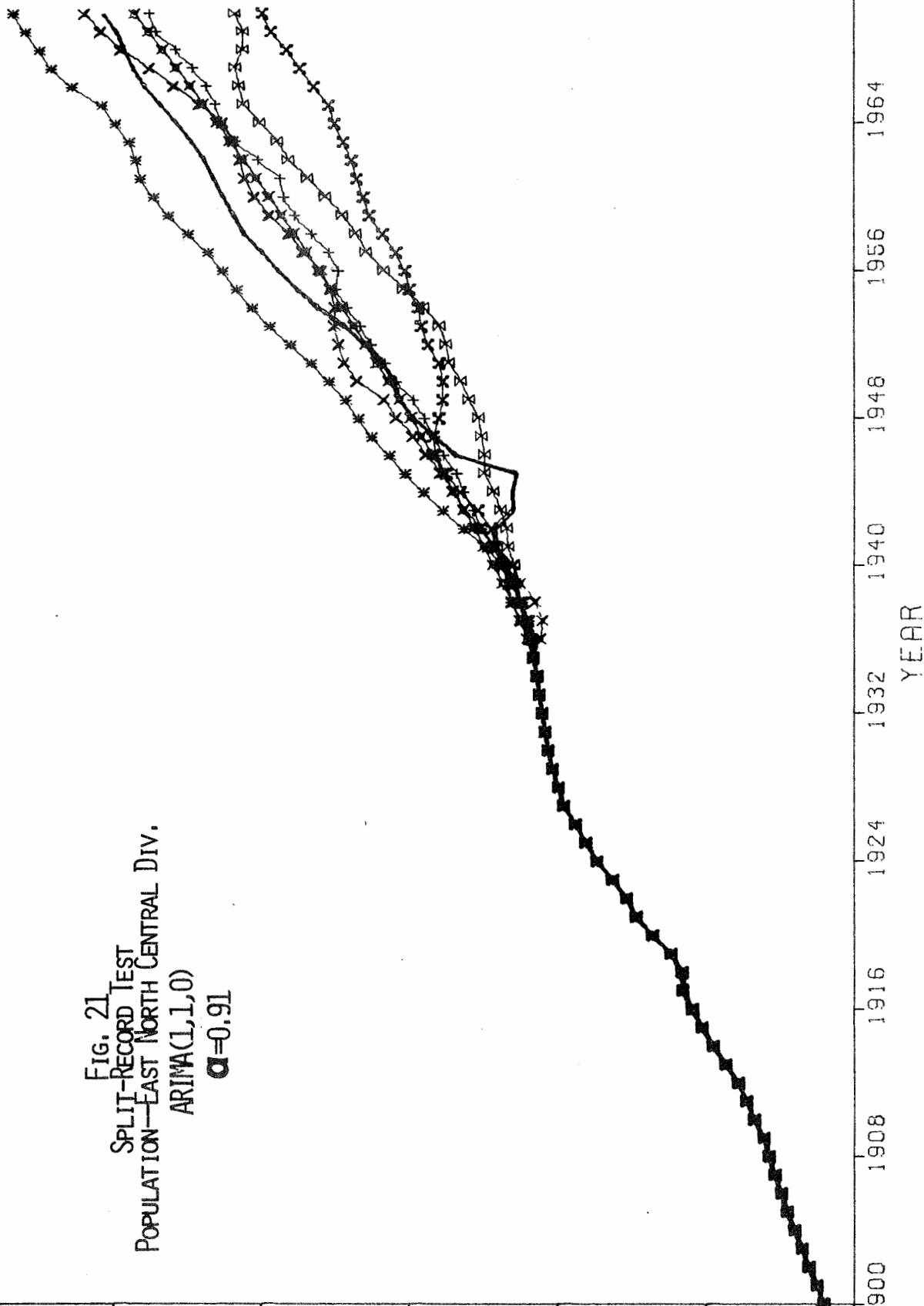


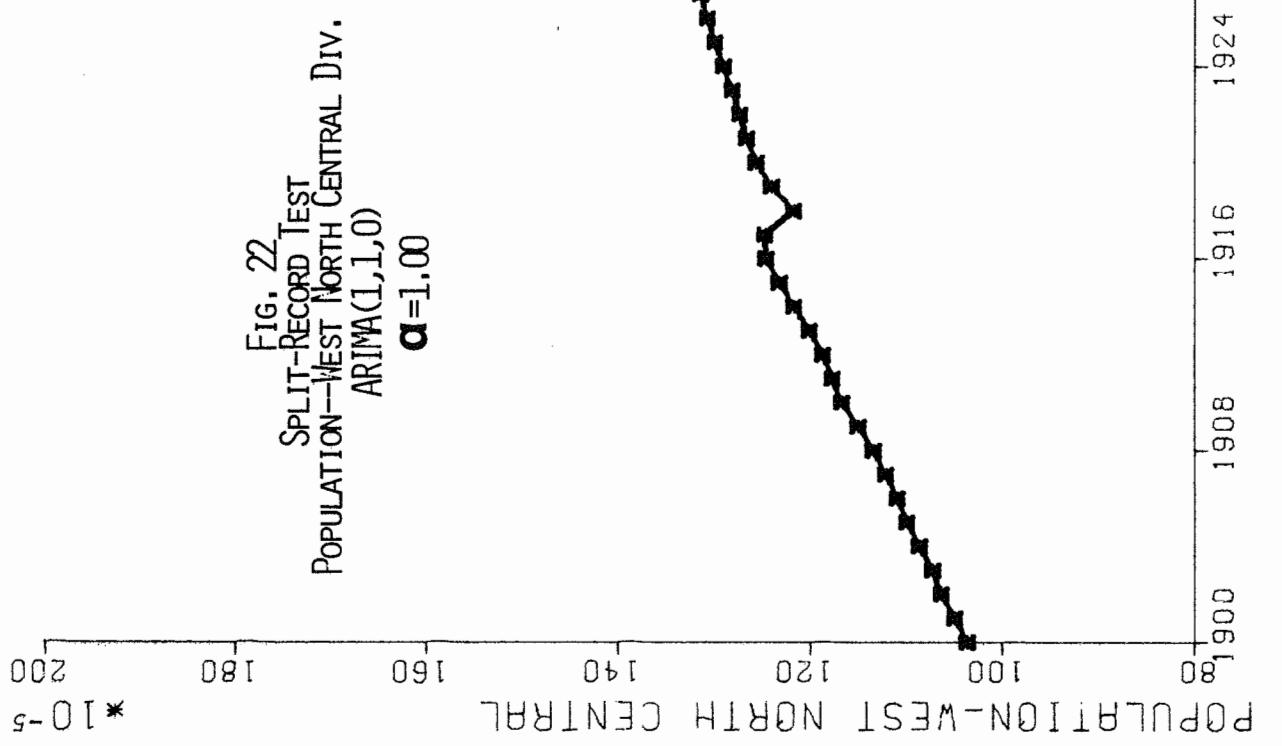
Fig. 20
SPLIT-RECORD TEST
POPULATION—MIDDLE ATLANTIC DIV.
ARIMA(1,1,0)
 $\alpha=0.96$

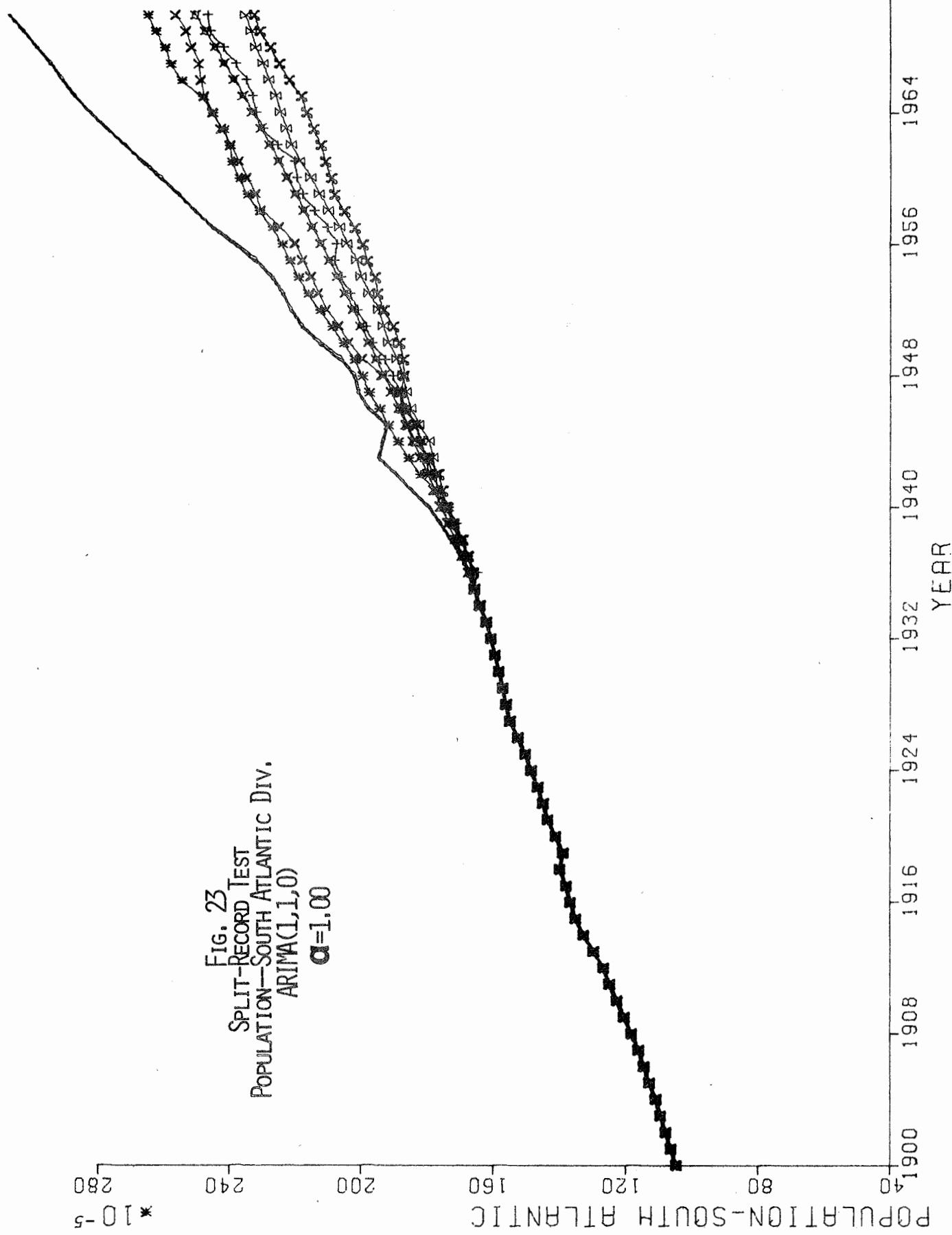


POPULATION-EAST NORTH CENTRAL
* 10⁻⁵

FIG. 21
SPLIT-RECORD TEST
POPULATION-EAST NORTH CENTRAL DIV.
ARIMA(1,1,0)
 $\alpha=0.91$







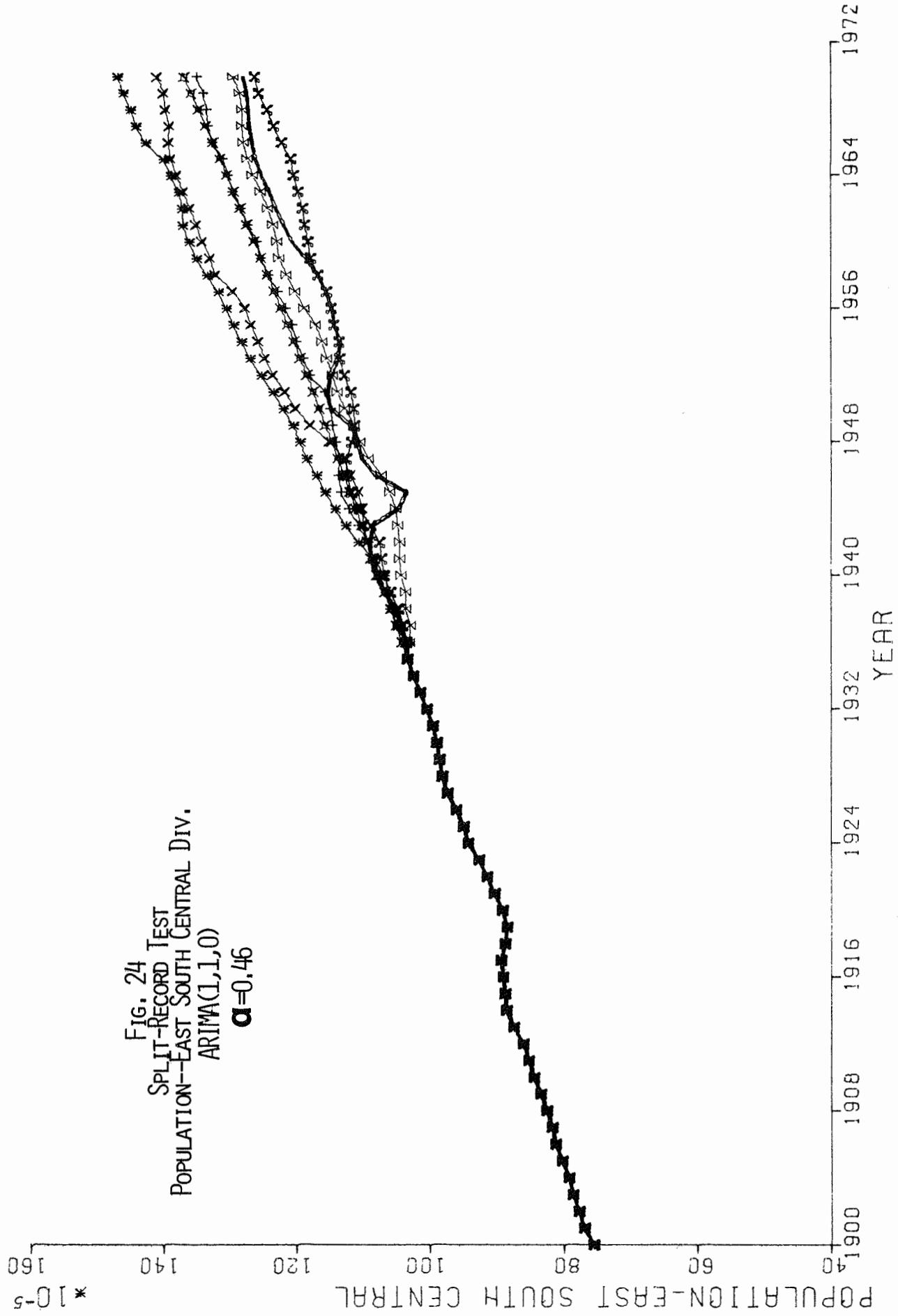
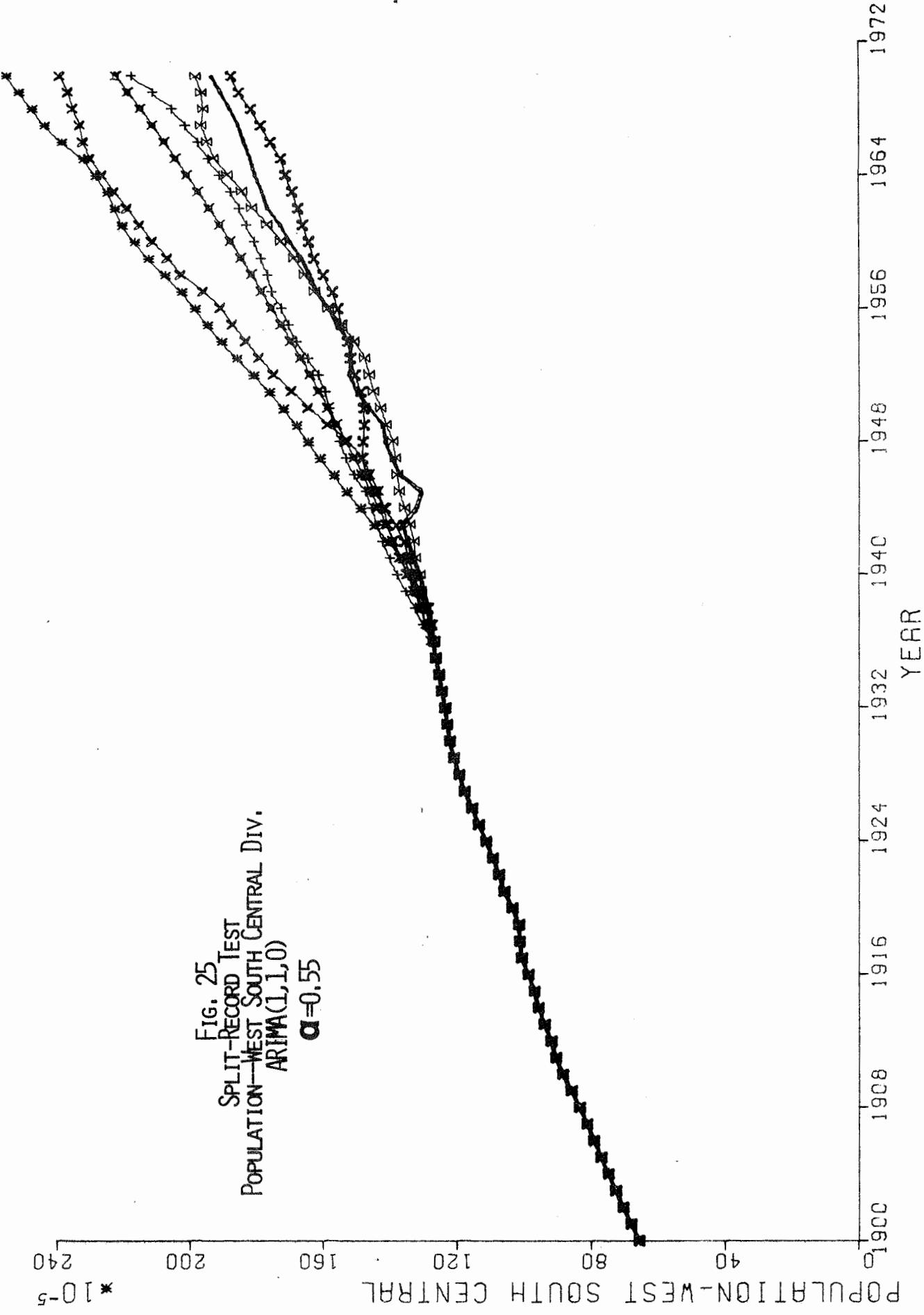
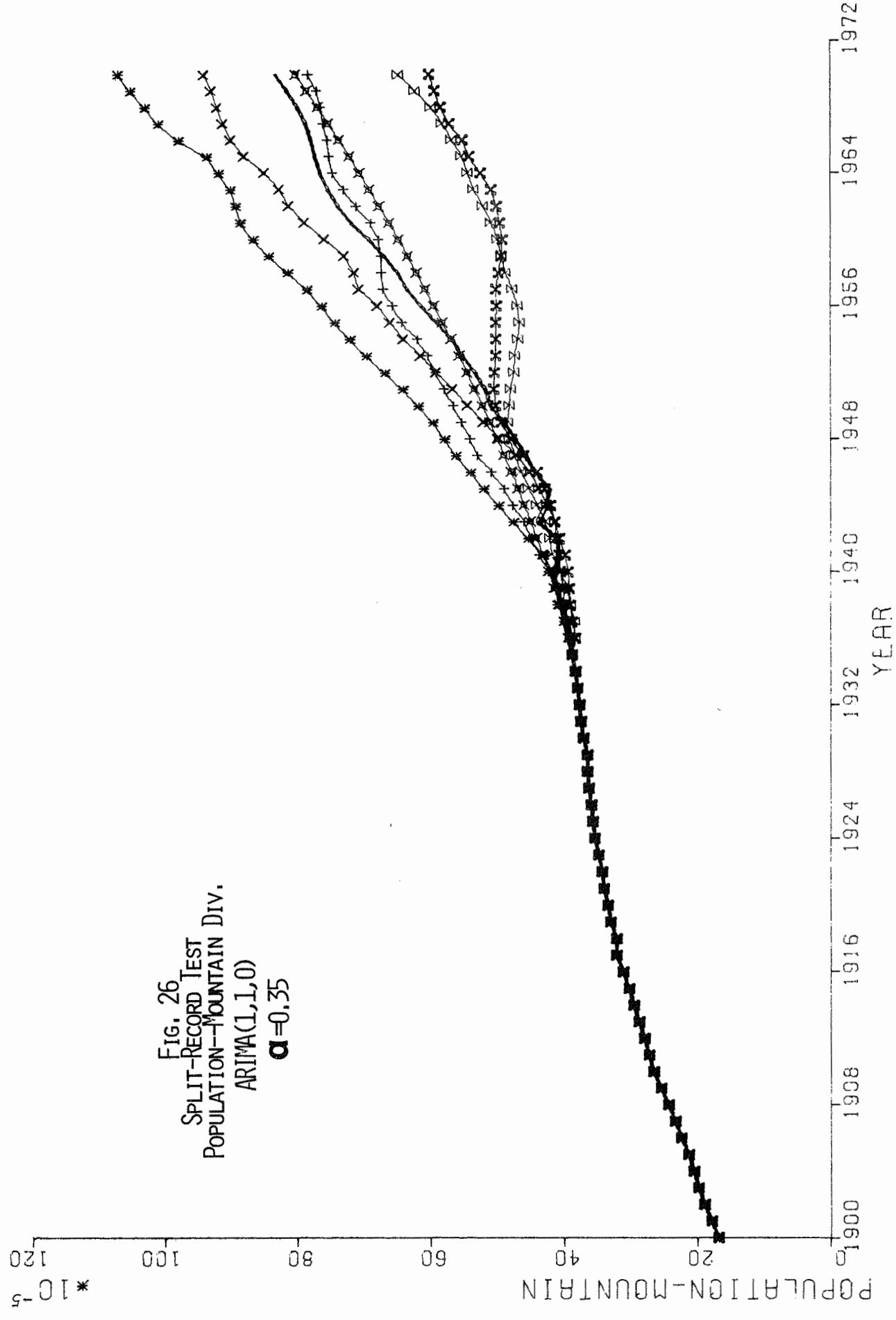


Fig. 25
SPLIT-RECORD TEST
POPULATION—WEST SOUTH CENTRAL DIV.
ARIMA(1,1,0)
 $\alpha = 0.55$





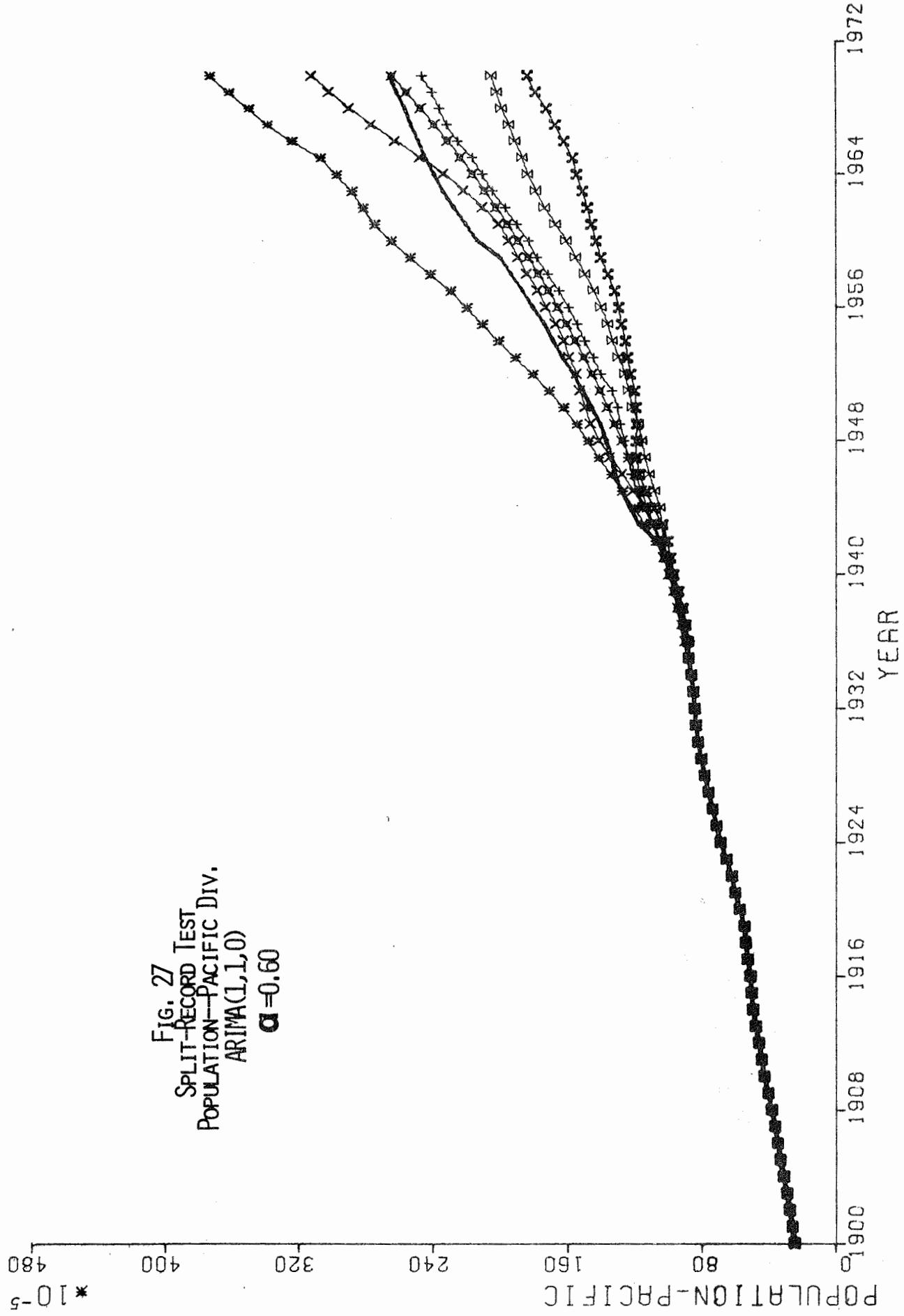


FIG. 28
SPLIT-RECORD TEST
POPULATION—NEW ENGLAND DIV.
MULTIVARIATE PROJECTION
PROJECTIONS

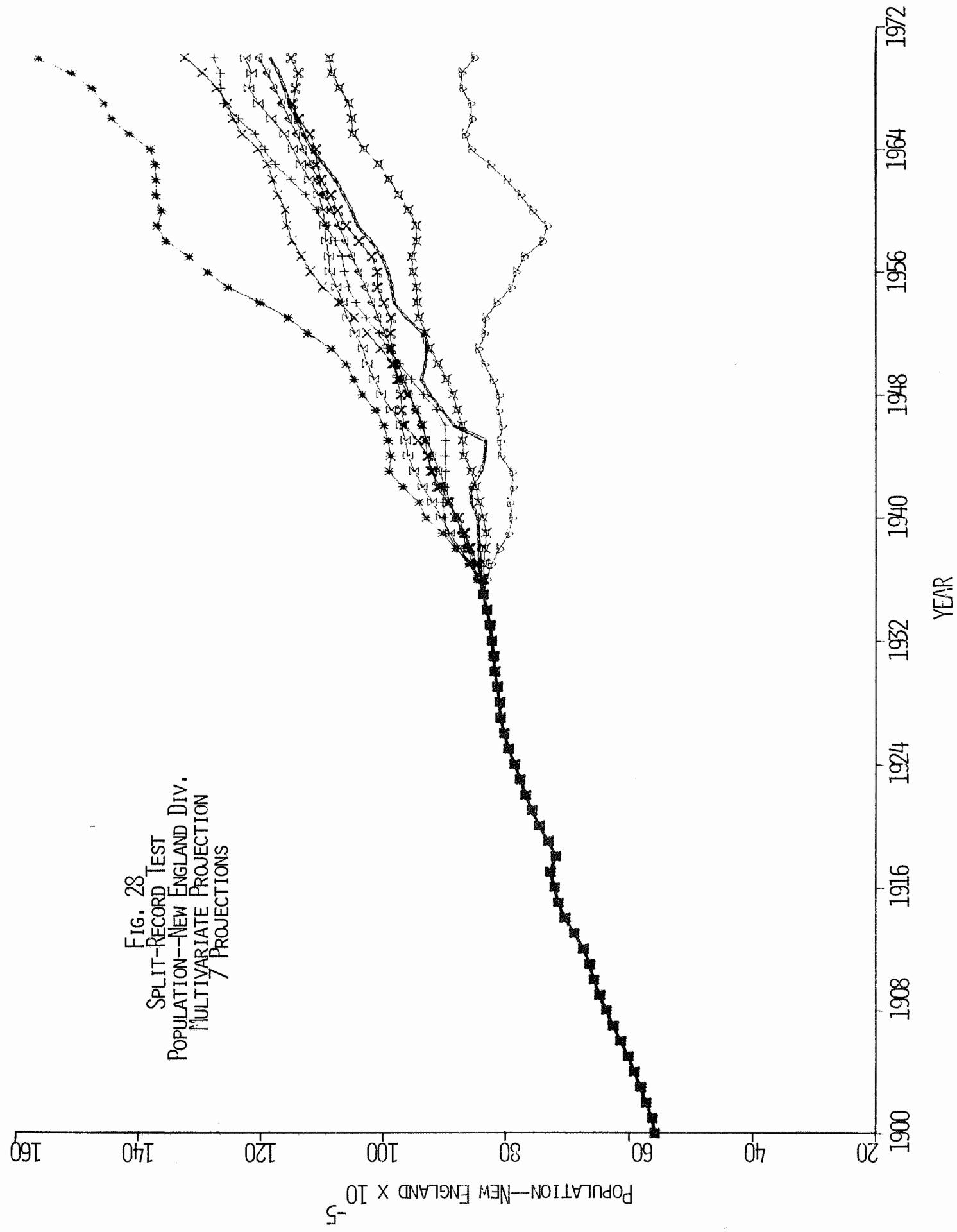


FIG. 29
SPLIT-RECORD TEST
POPULATION-MIDDLE ATLANTIC DIV.
MULTIVARIATE PROJECTION
PROJECTIONS

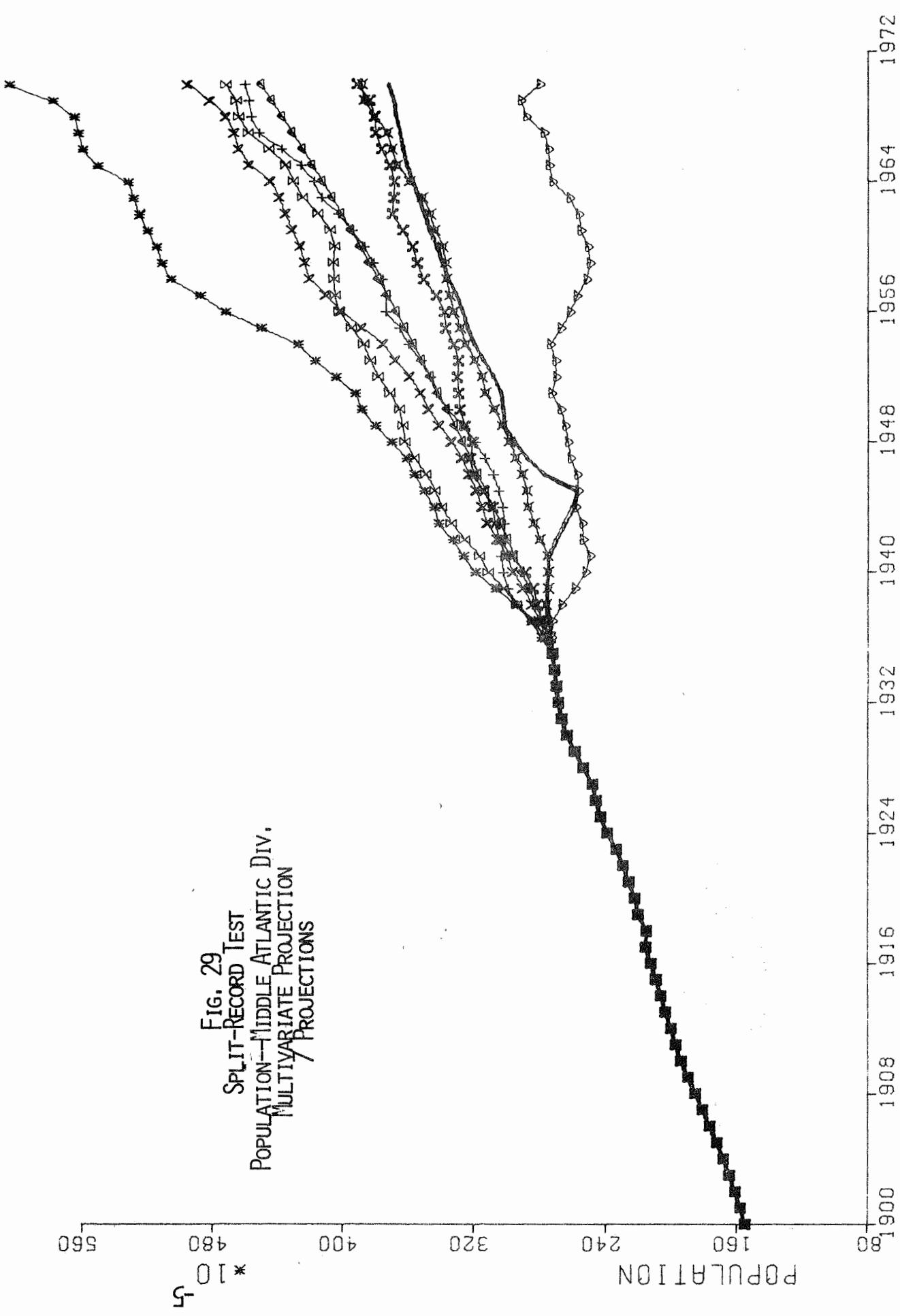
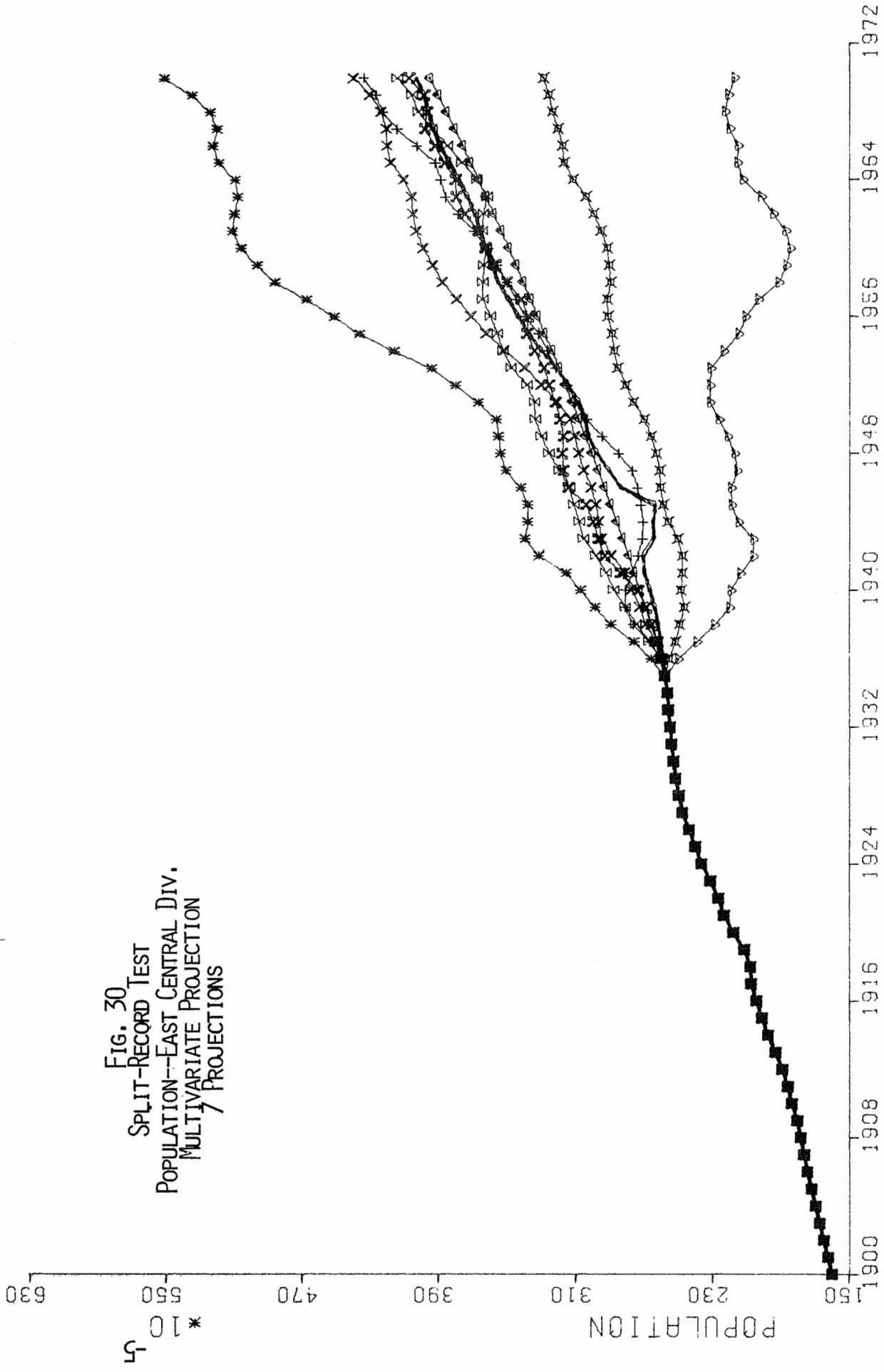


FIG. 30
SPLIT-RECORD TEST
POPULATION--EAST CENTRAL DIV.
MULTIVARIATE PROJECTION
PROJECTIONS



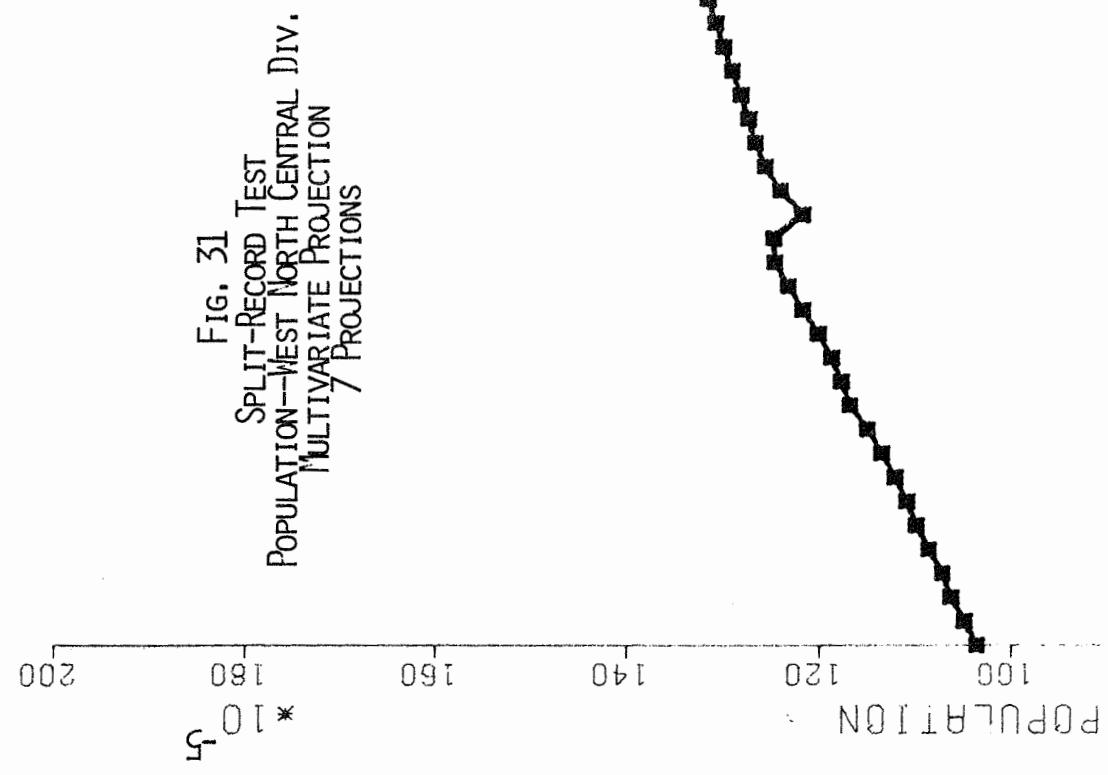
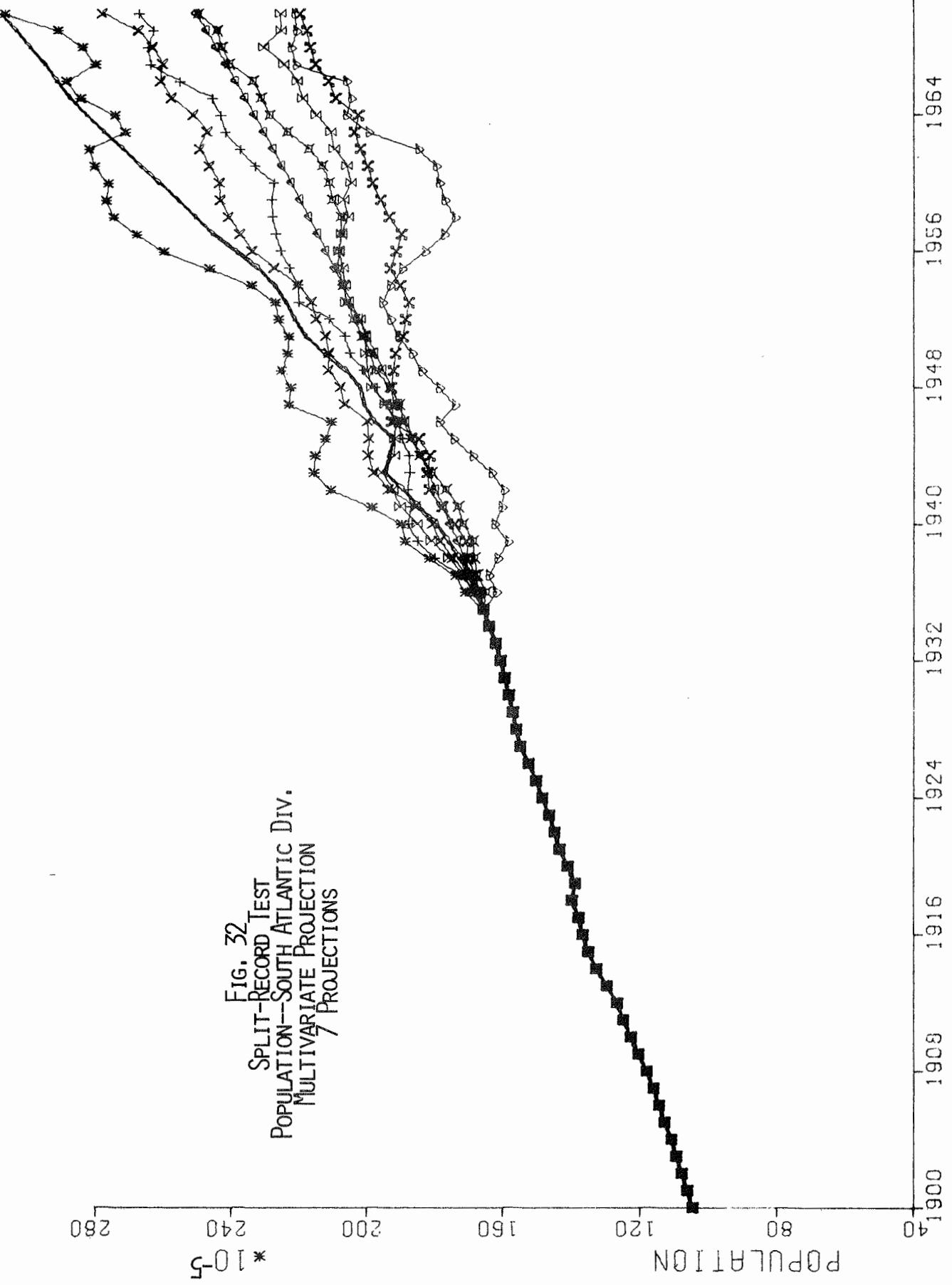


FIG. 31
SPLIT-RECORD TEST
POPULATION—WEST NORTH CENTRAL DIV.
MULTIVARIATE PROJECTION
PROJECTIONS



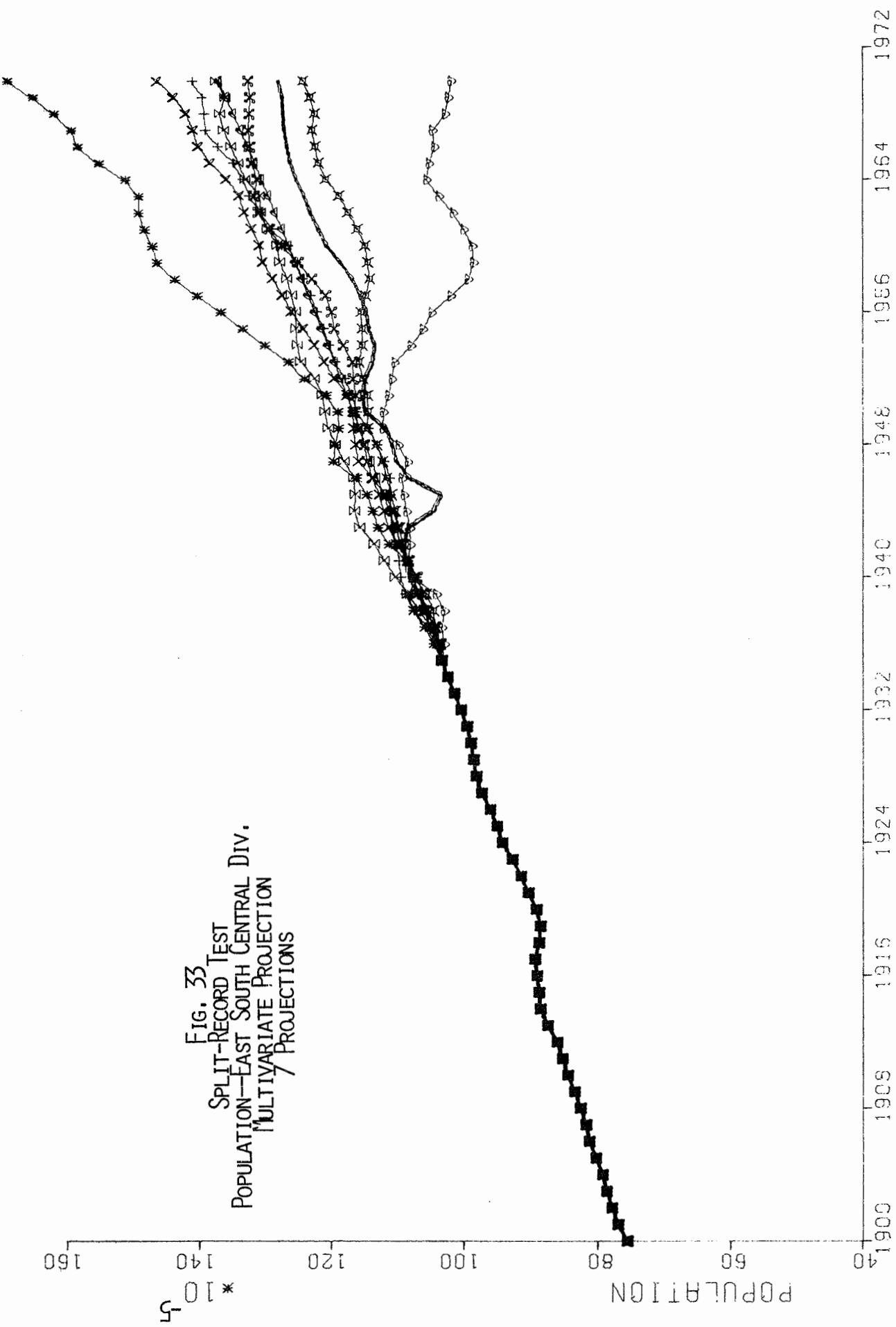


FIG. 33
SPLIT-RECORD TEST
POPULATION—EAST SOUTH CENTRAL DIV.
MULTIVARIATE PROJECTION
PROJECTIONS

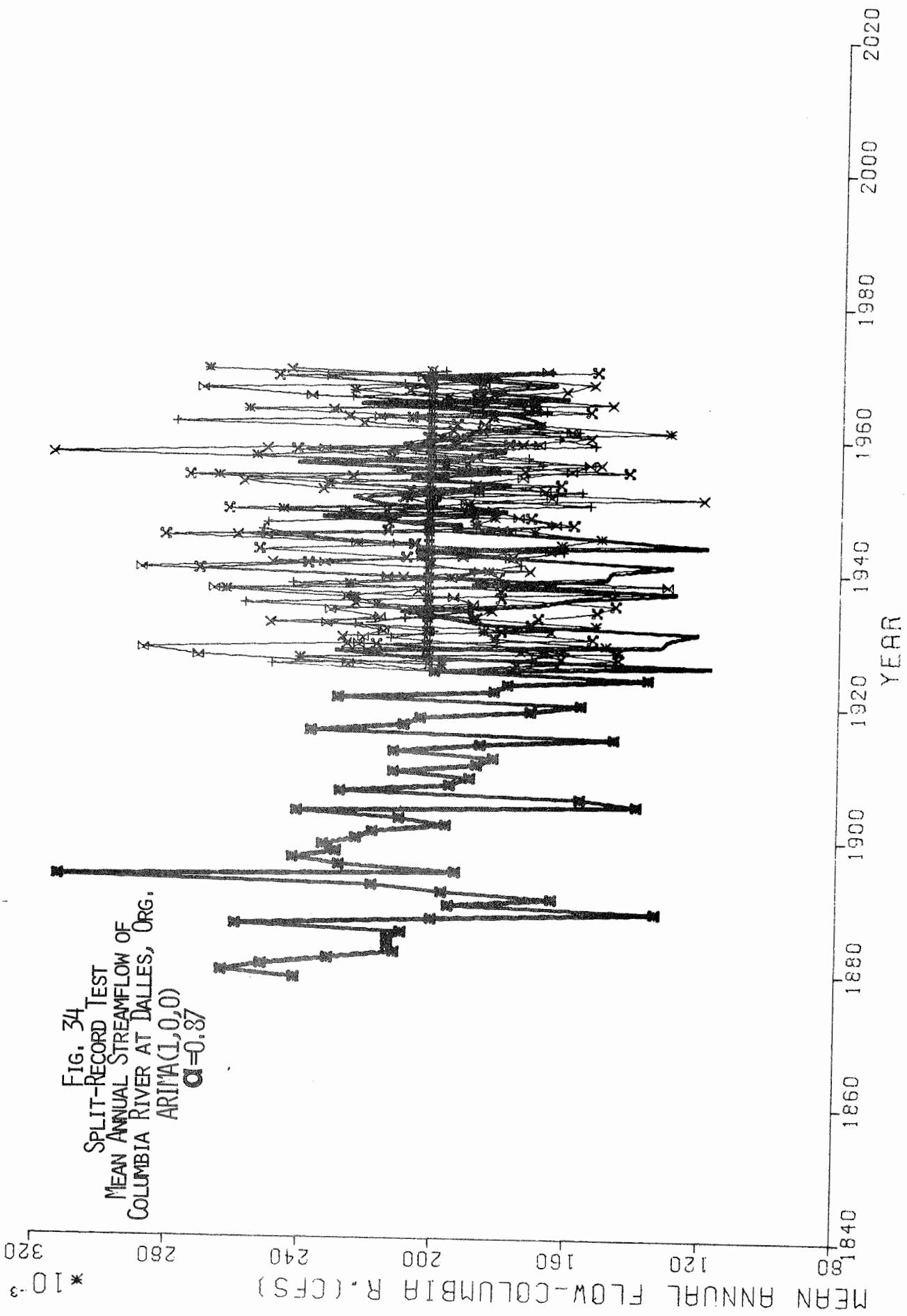
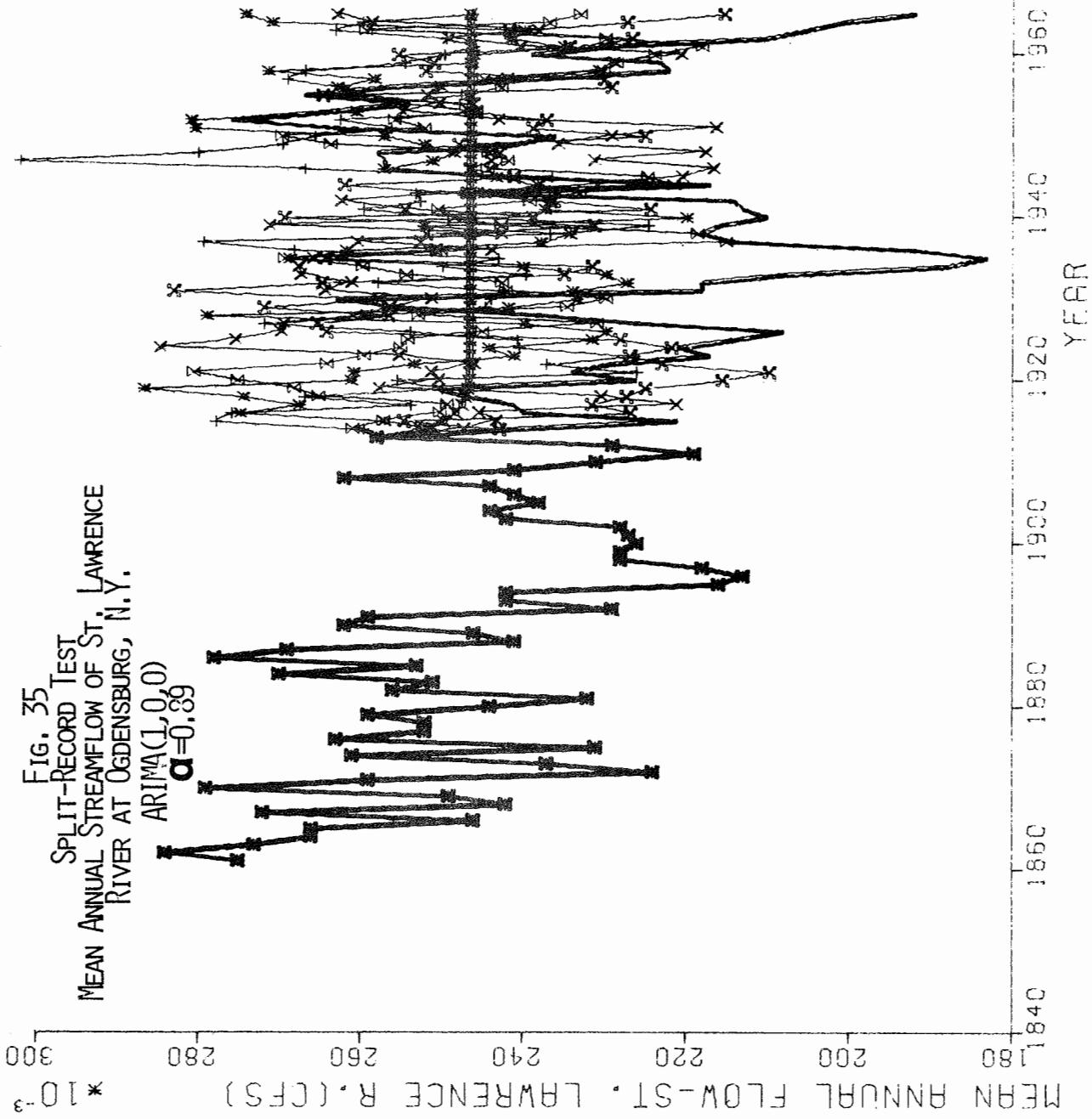


FIG. 35
SPLIT-RECORD TEST
MEAN ANNUAL STREAMFLOW OF ST. LAWRENCE
RIVER AT OGDENSBURG, N.Y.

$\text{ARIMA}(1,0,0)$
 $\alpha = 0.89$



MEAN ANNUAL FLOW-MISSISSIPPI R. (CFS) * 10^{-3}

FIG. 36
SPLIT-RECORD TEST
MEAN ANNUAL STREAMFLOW OF
MISSISSIPPI RIVER AT ST. LOUIS, Mo.
ARIMA(1,0,0)
 $\alpha = 0.33$

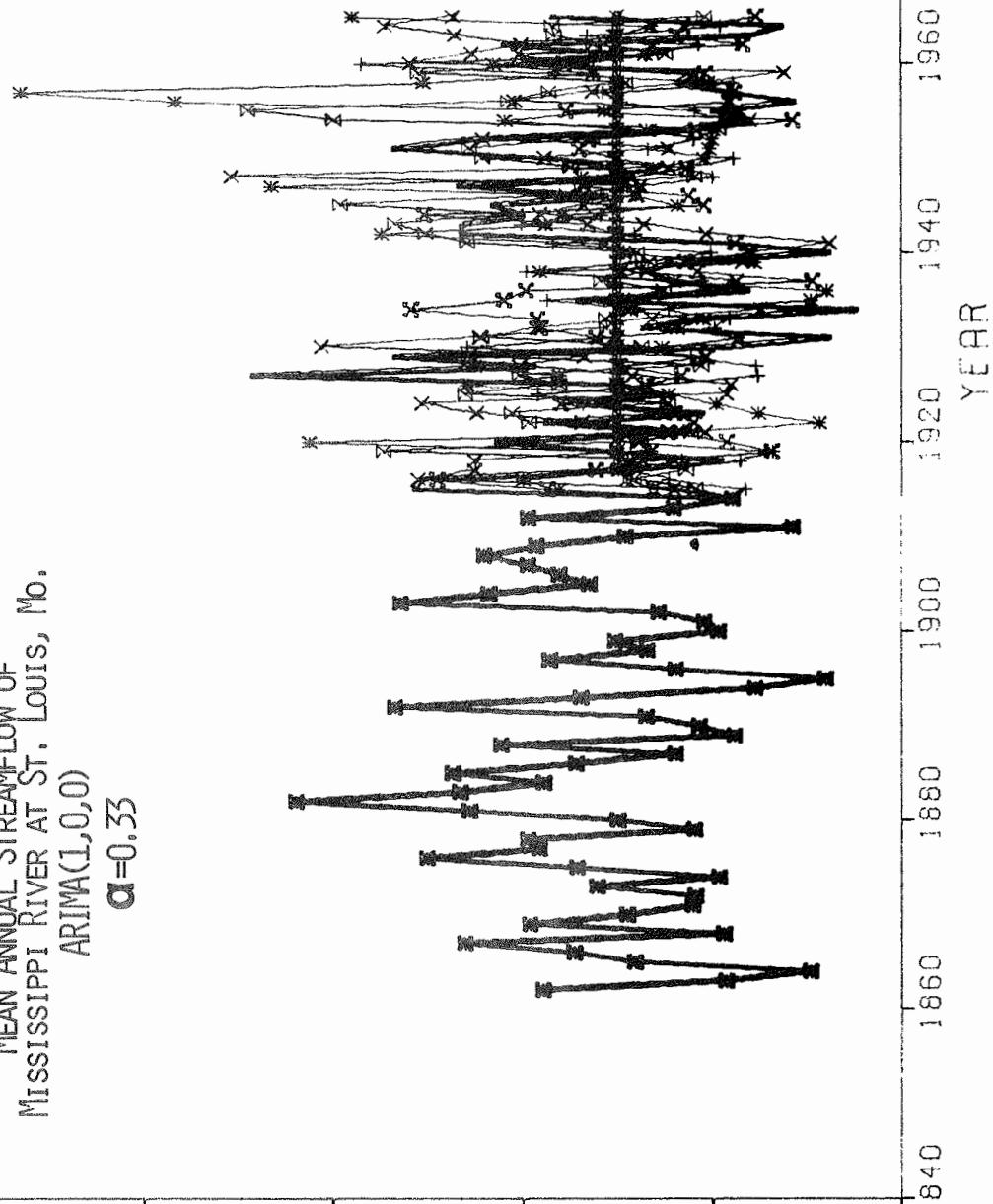
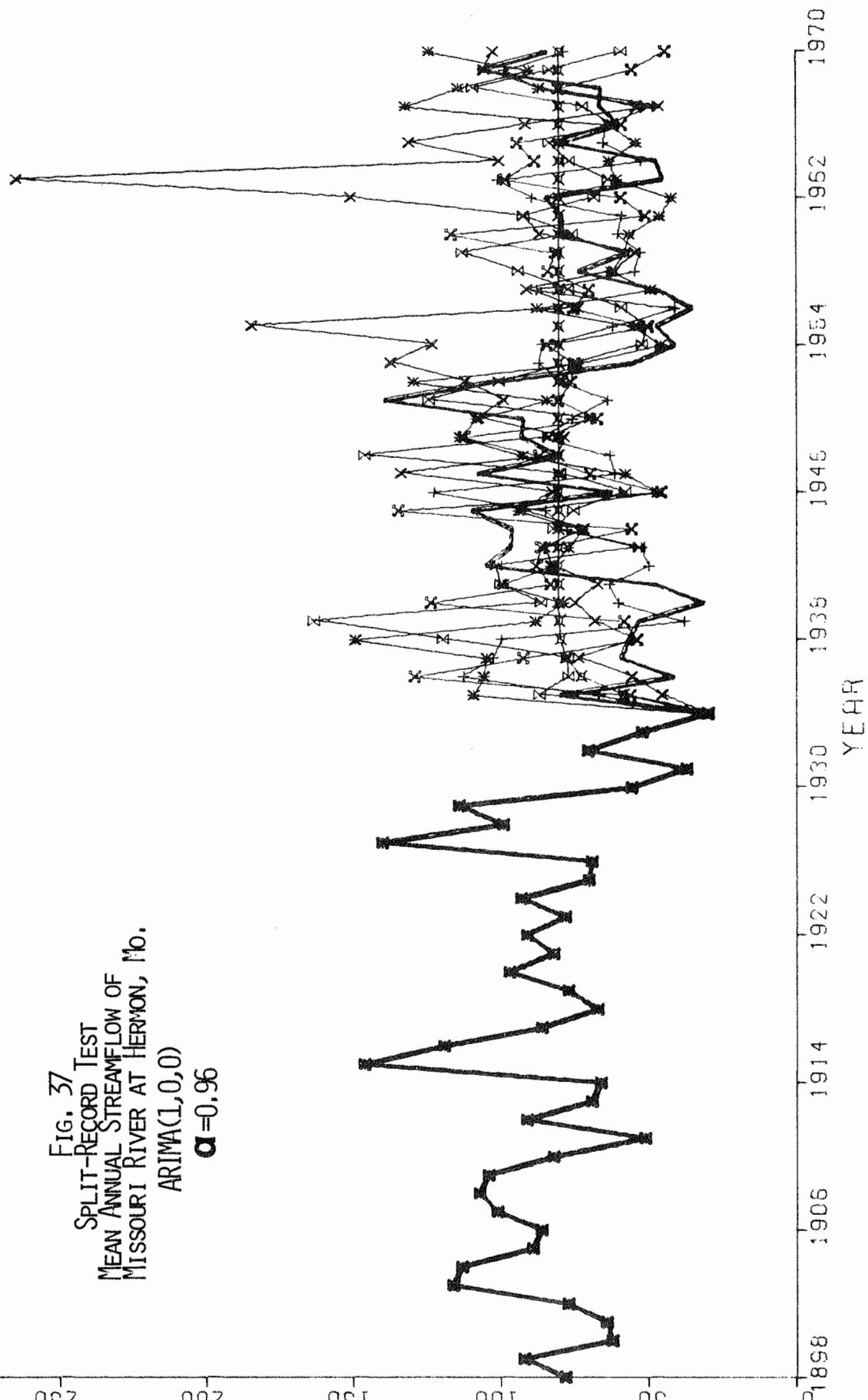
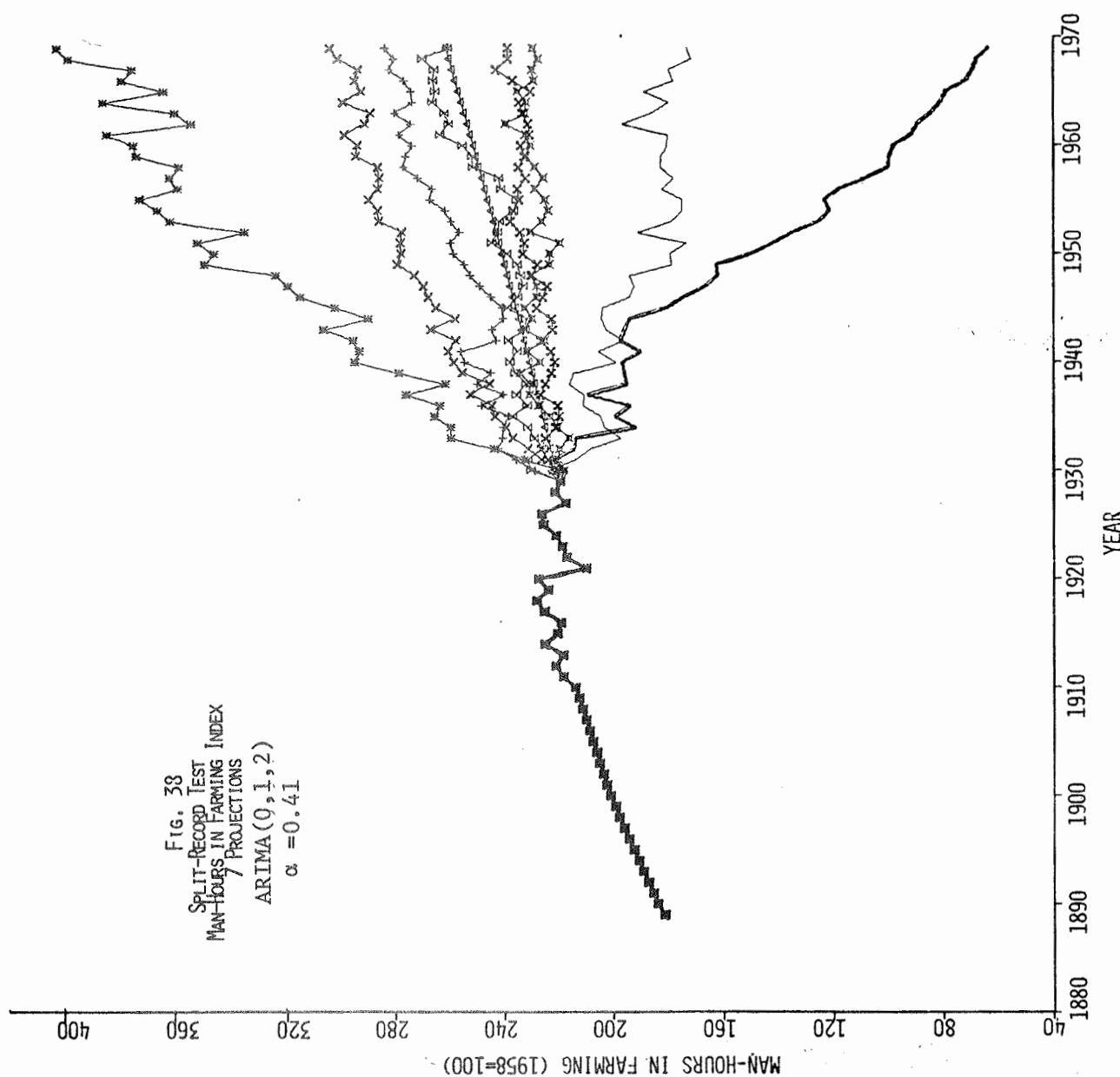
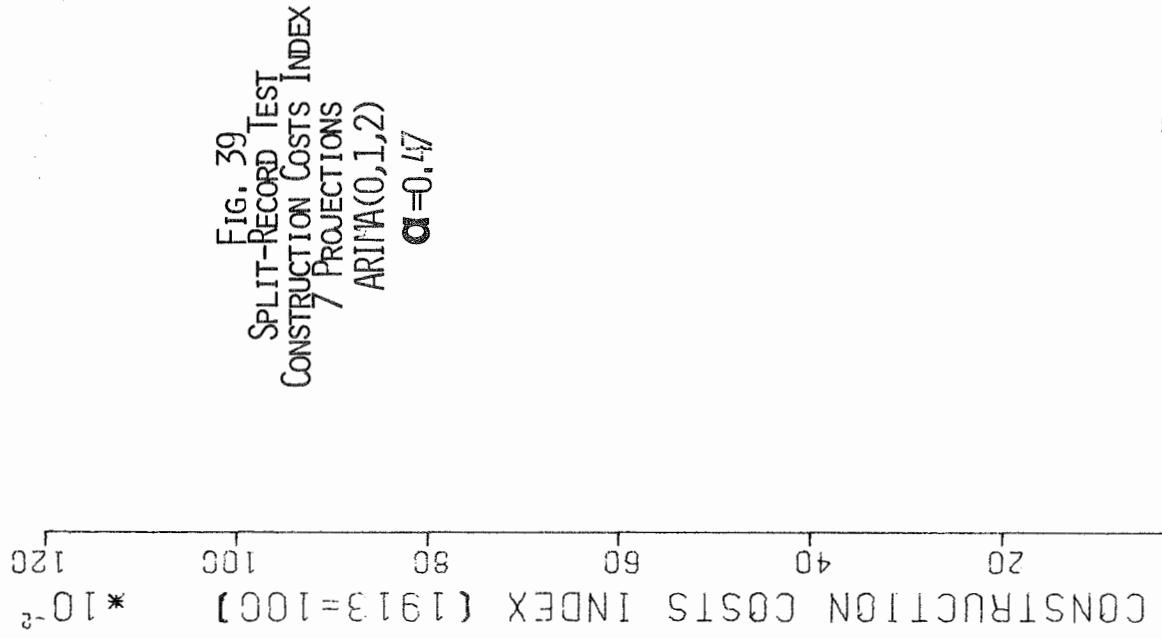


FIG. 37
SPLIT-RECORD TEST
MEAN ANNUAL STREAMFLOW OF
MISSOURI RIVER AT HERMON, Mo.
ARIMA(1,0,0)
 $\alpha = 0.96$







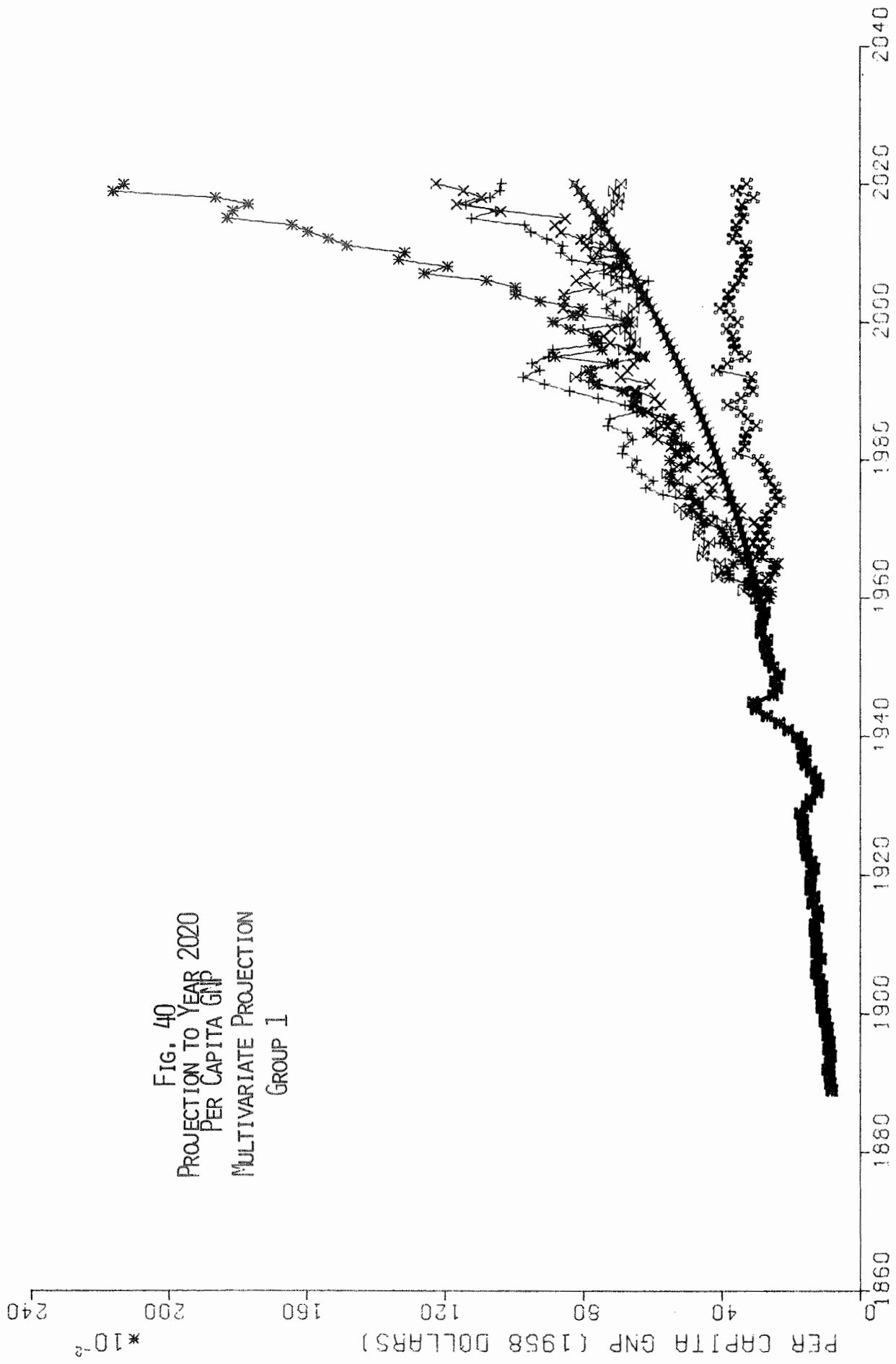
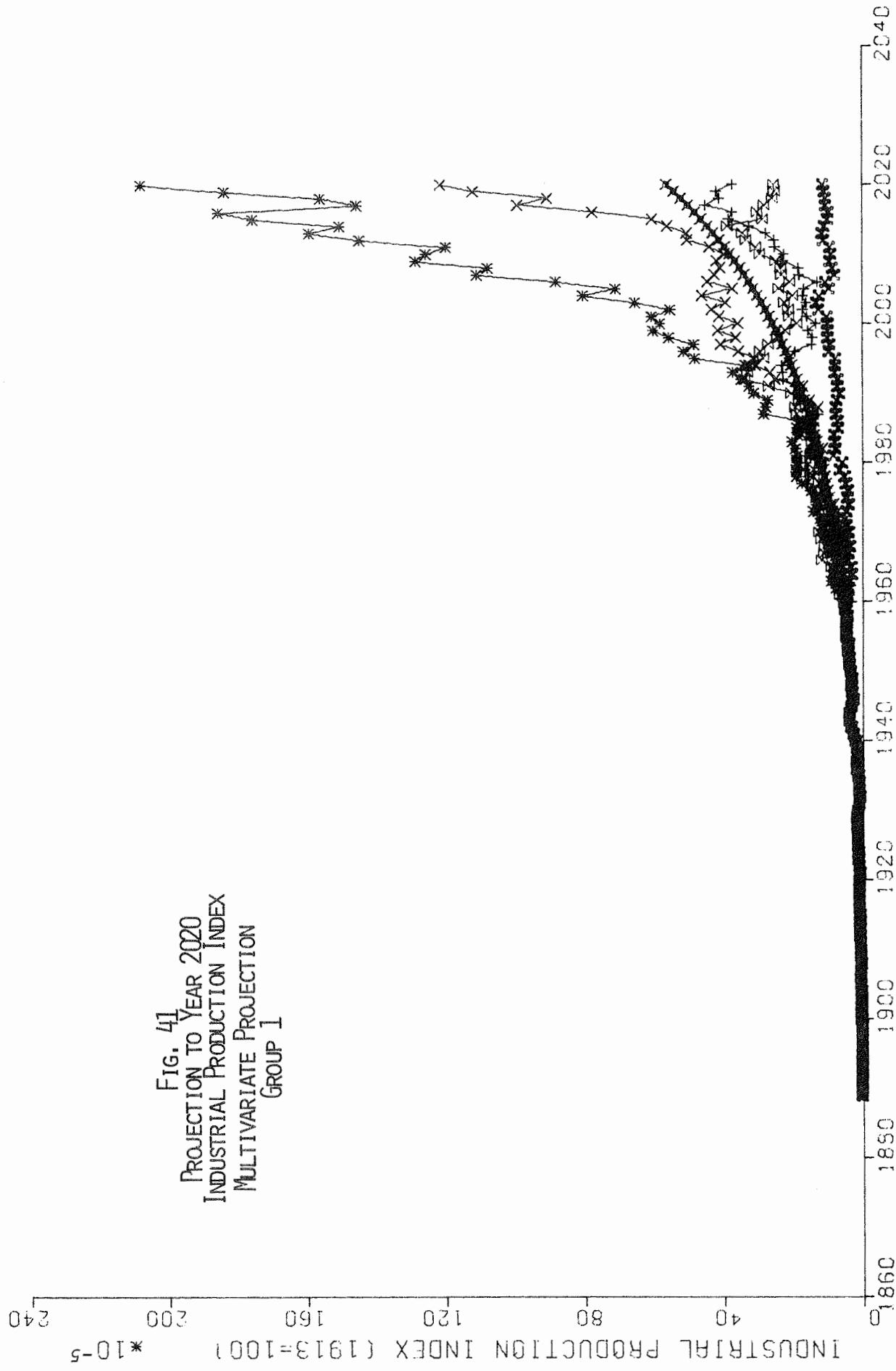
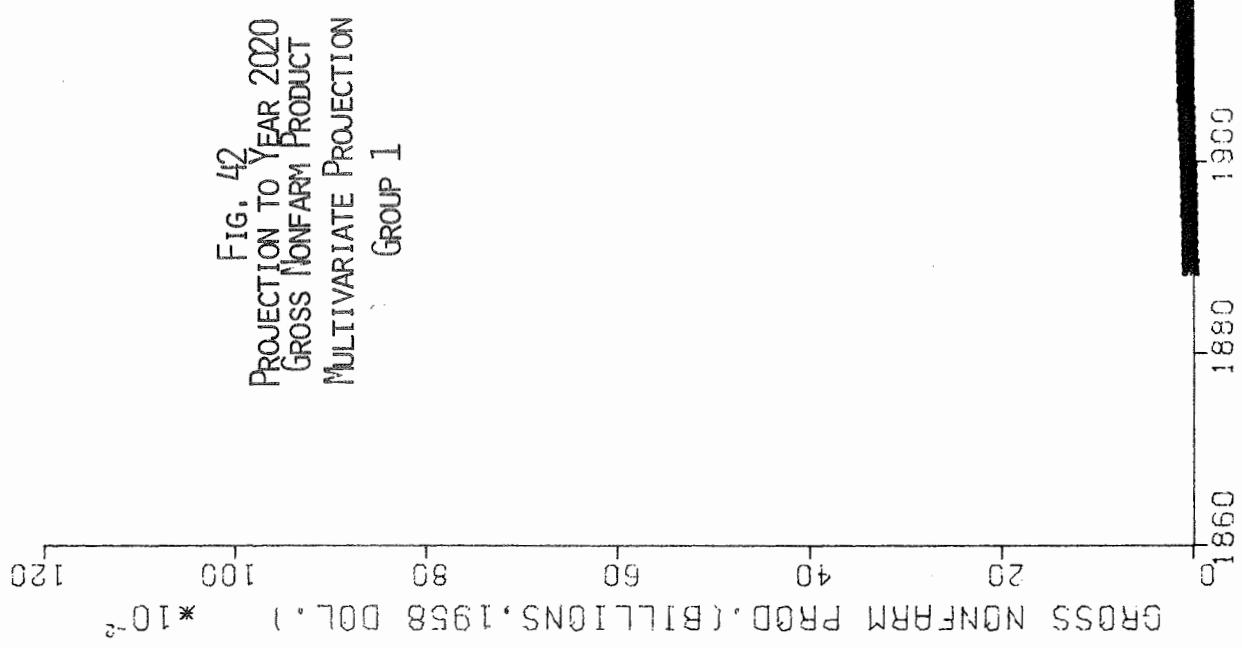
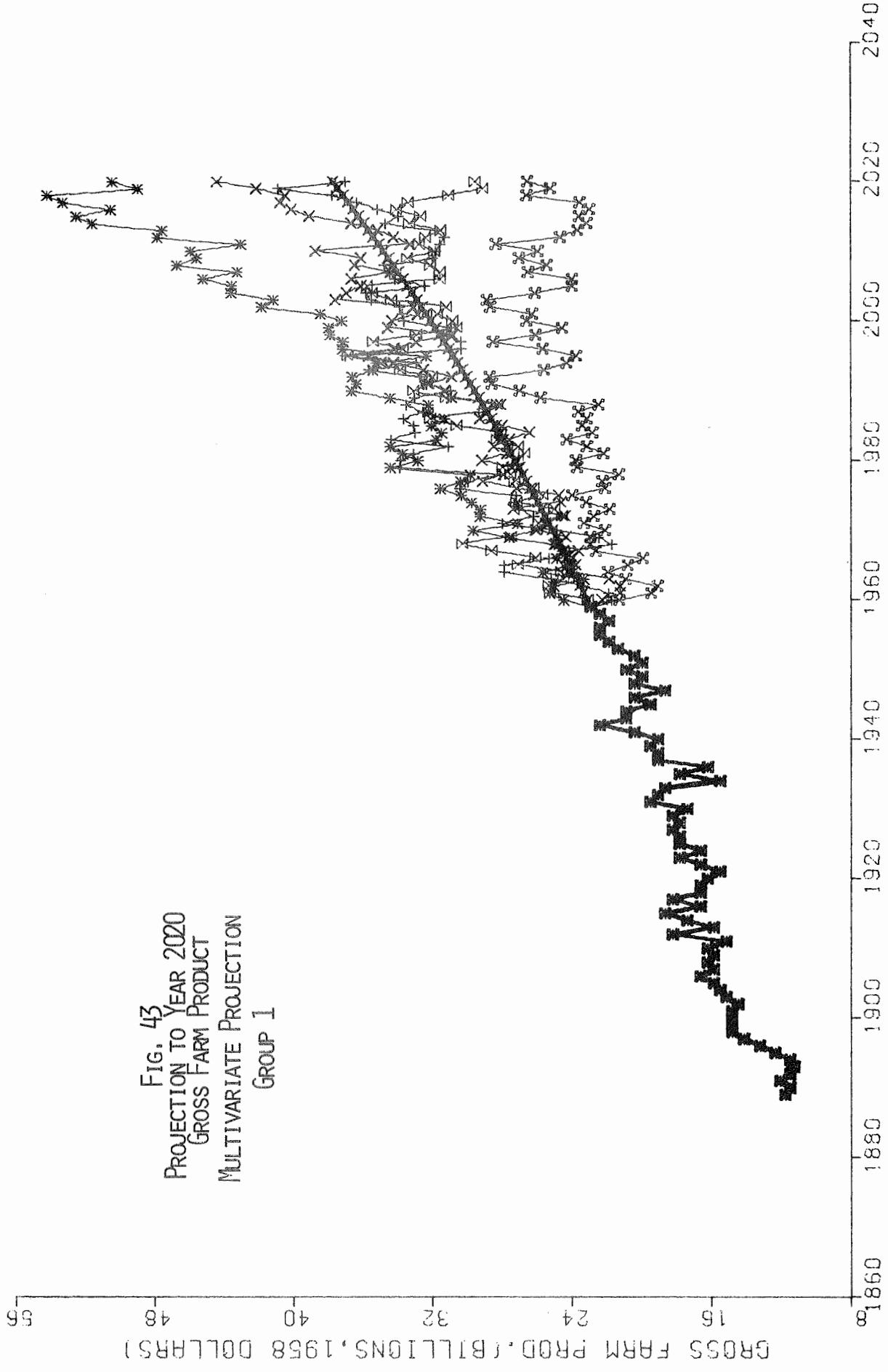
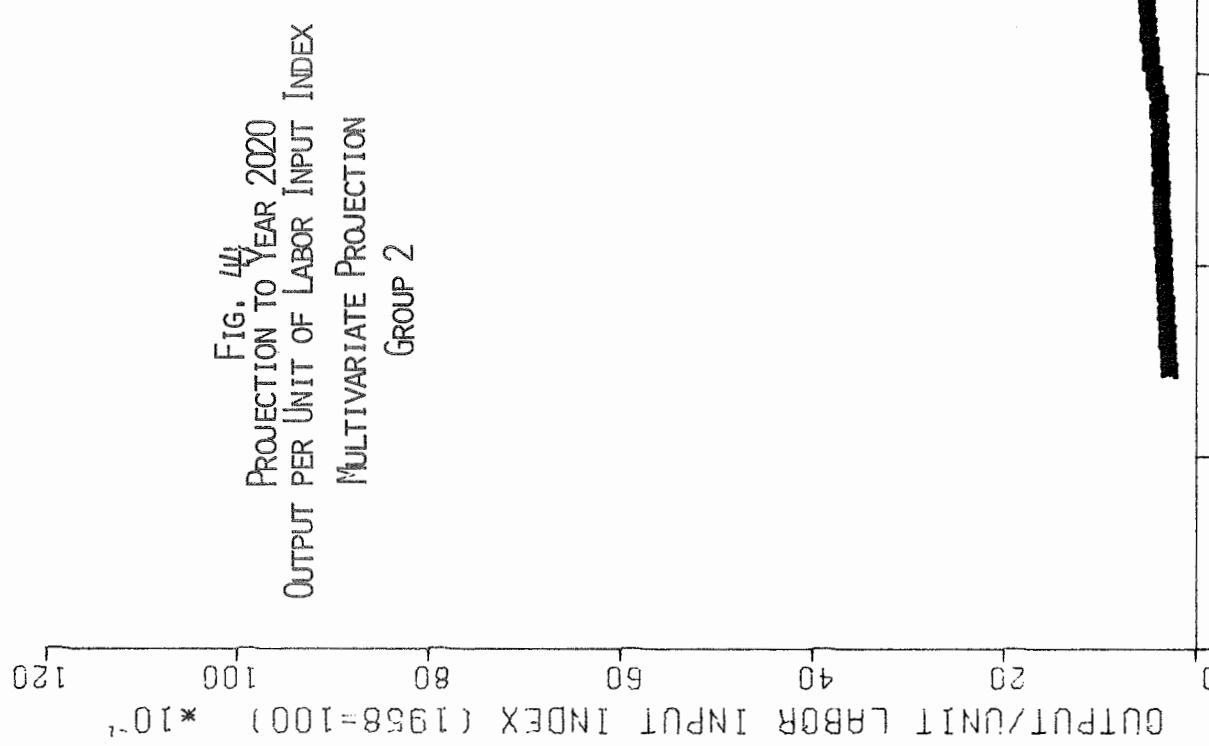


FIG. 41
PROJECTION TO YEAR 2020
INDUSTRIAL PRODUCTION INDEX
MULTIVARIATE PROJECTION
GROUP 1









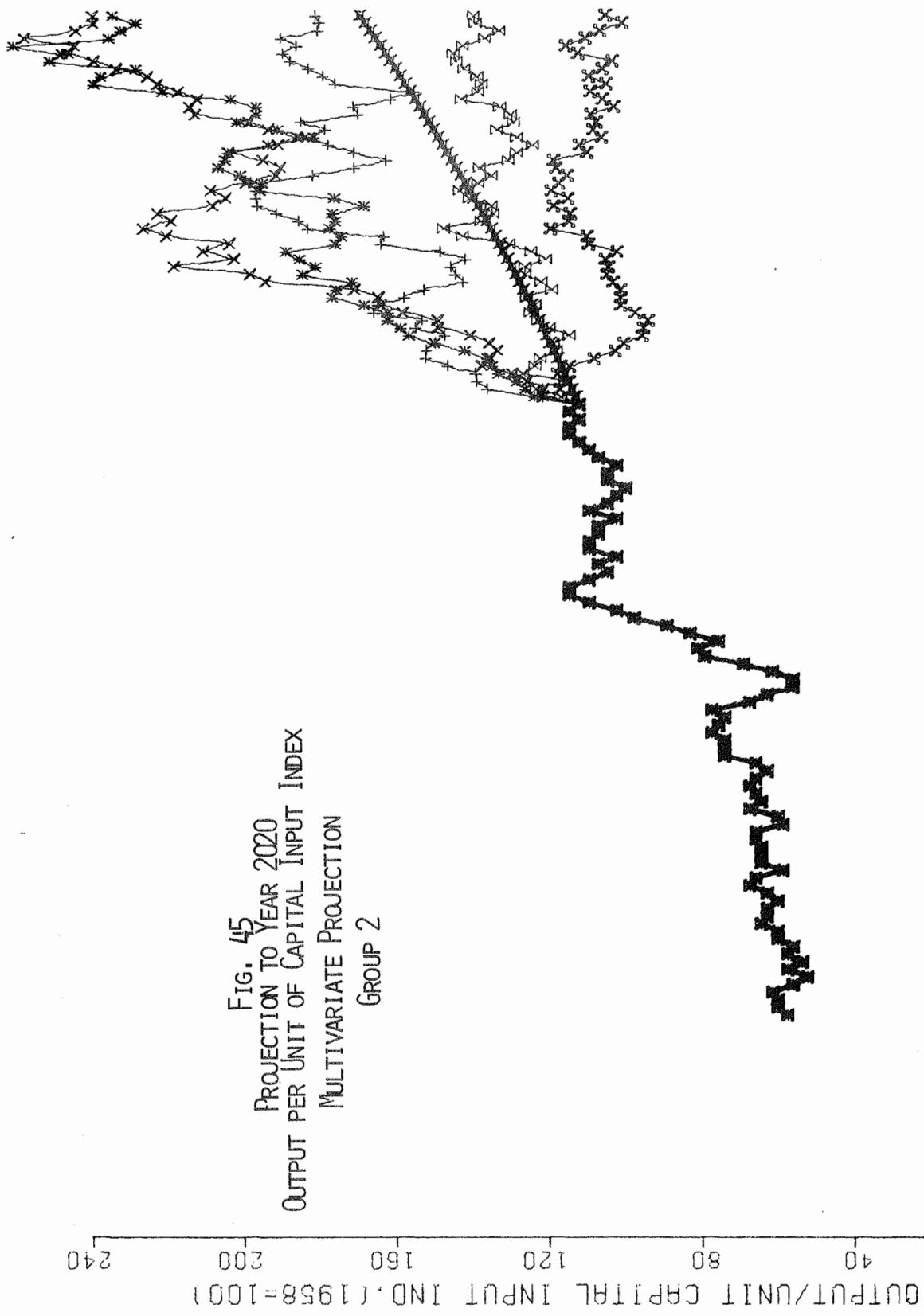


FIG. 46
PROJECTION TO YEAR 2020
WHOLESALE PRICE INDEX
MULTIVARIATE PROJECTION
GROUP 3

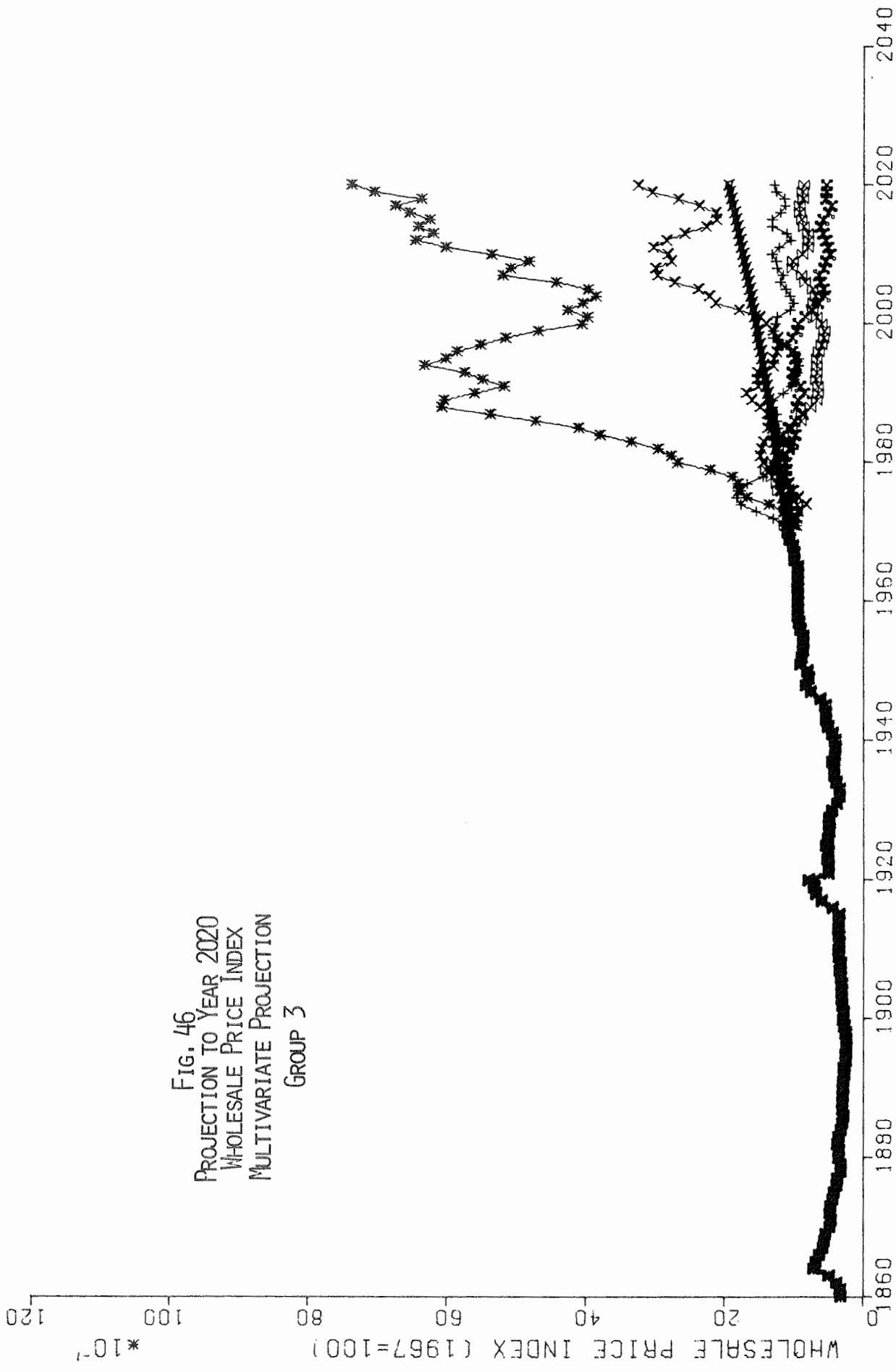
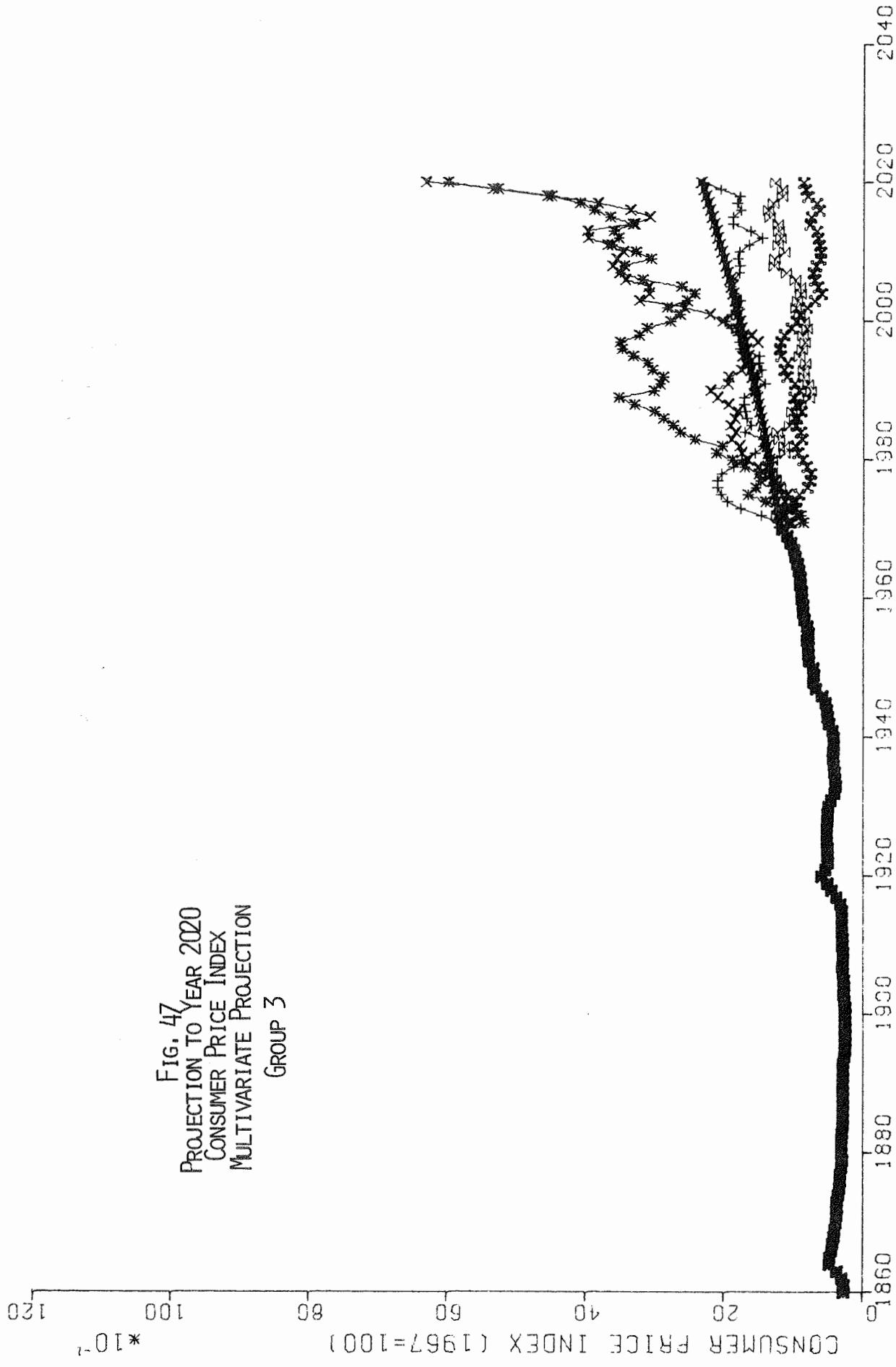


FIG. 47
PROJECTION TO YEAR 2020
CONSUMER PRICE INDEX
MULTIVARIATE PROJECTION
GROUP 3



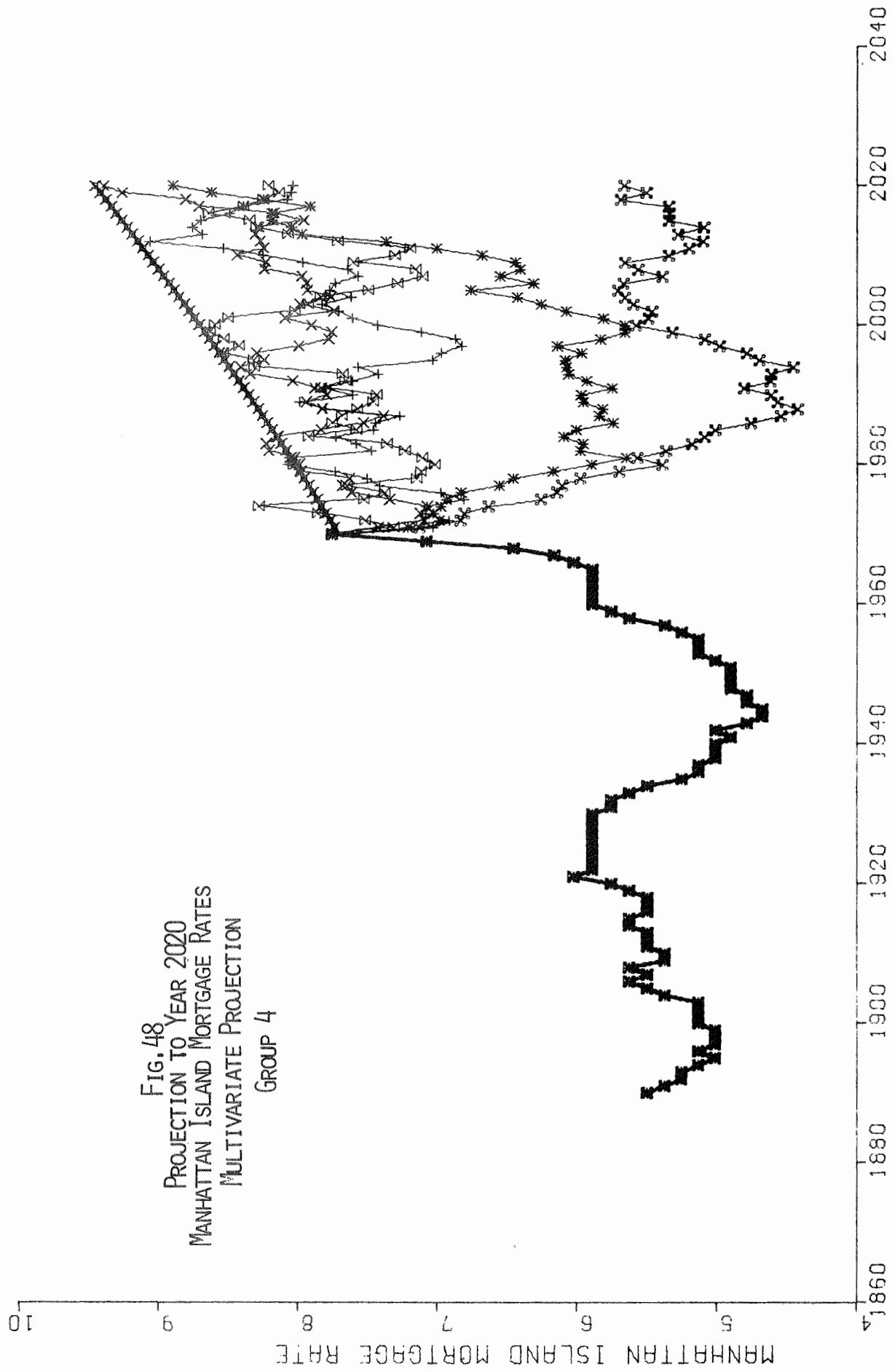


FIG. 49
PROJECTION TO YEAR 2020
Stock Exchange Call Loan Rates
Multivariate Projection
Group 4

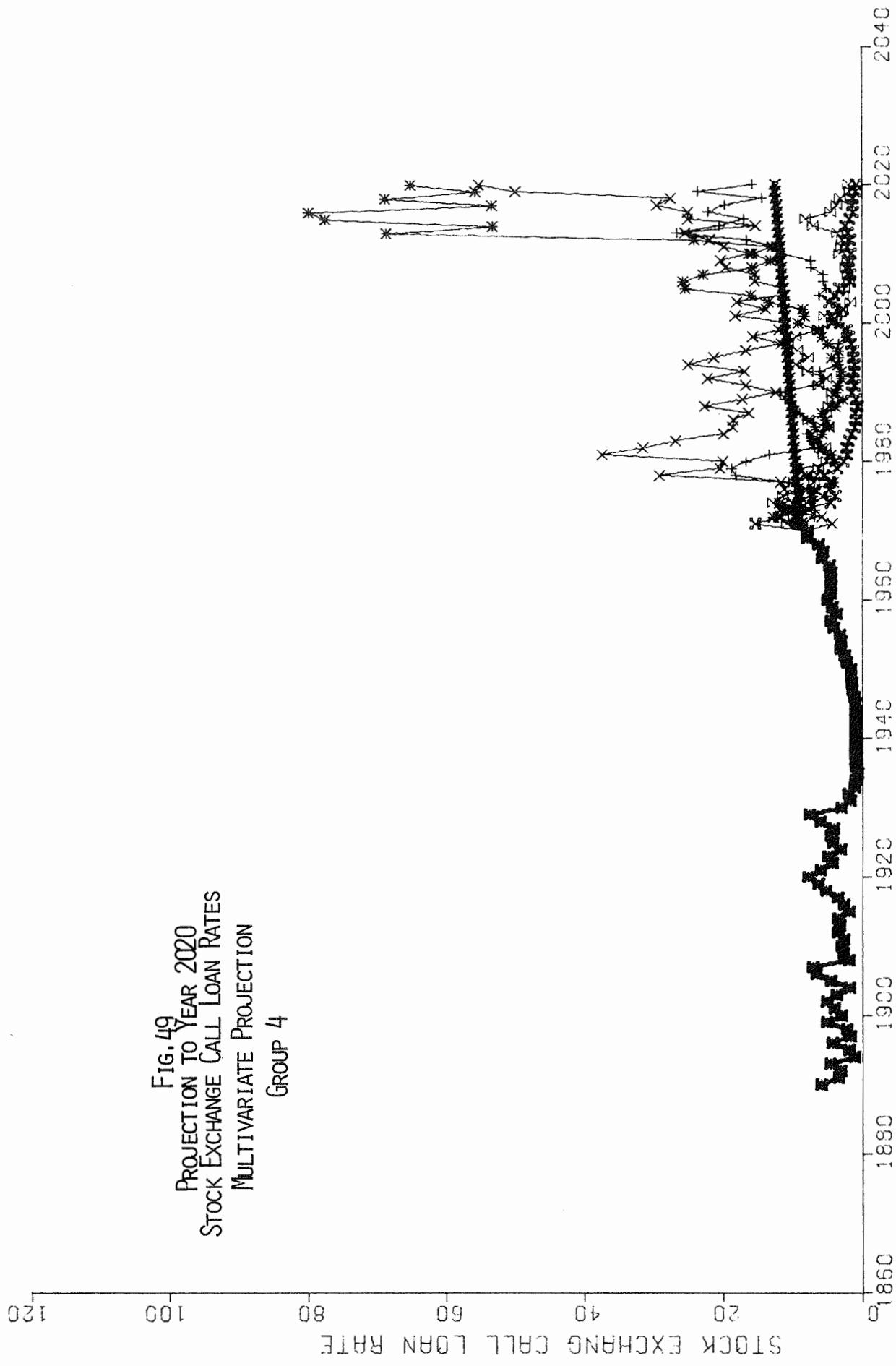
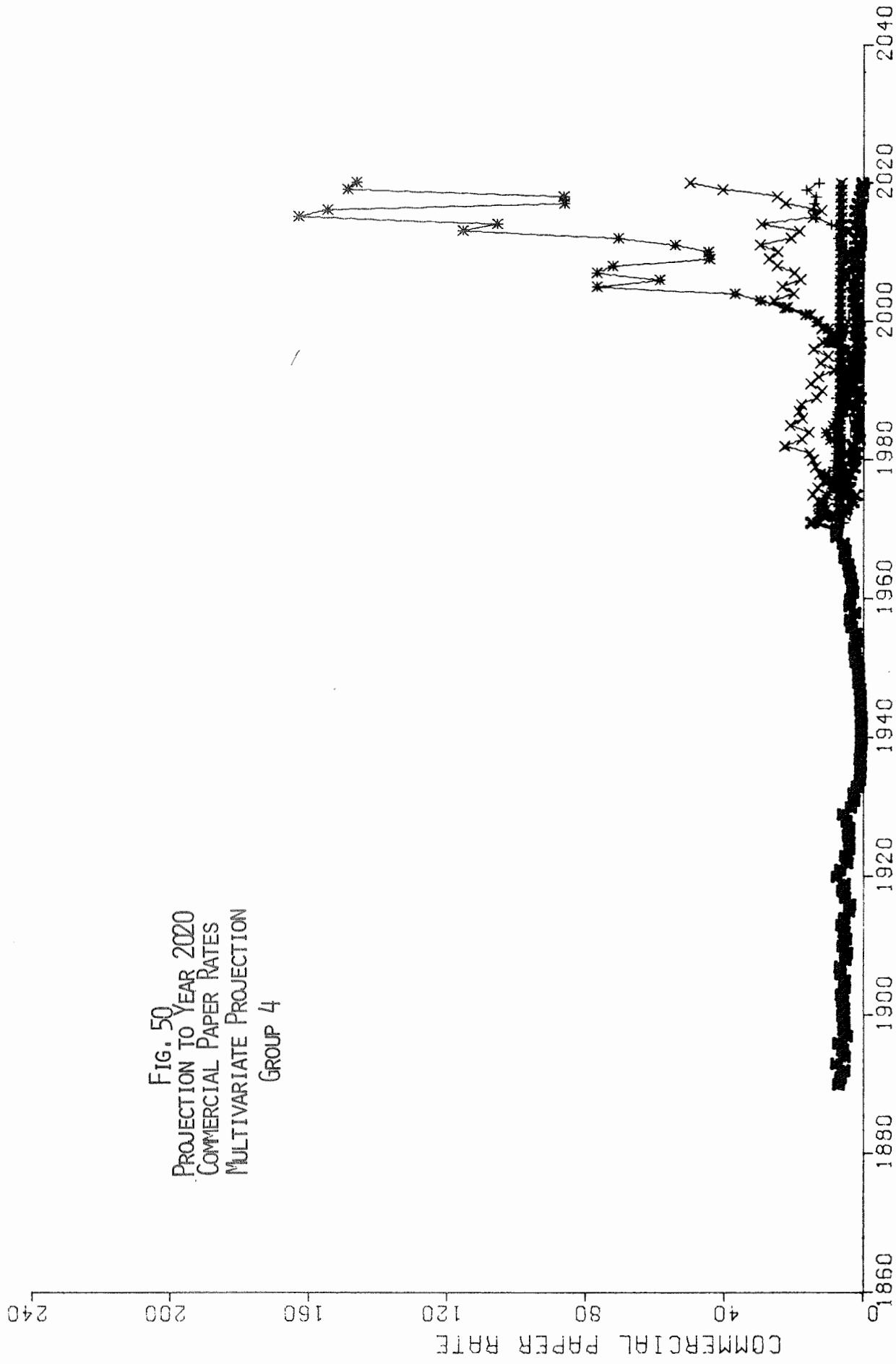
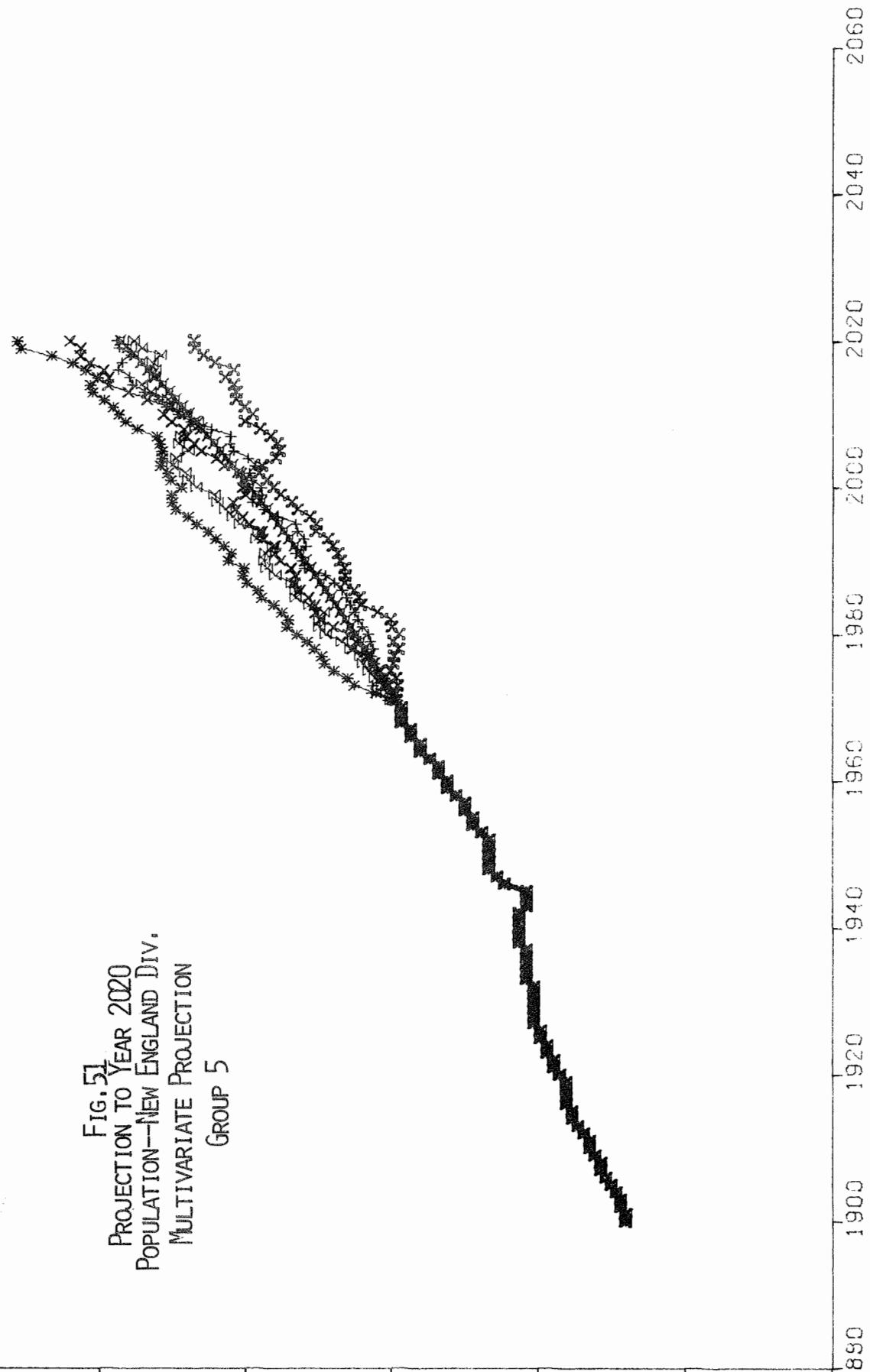
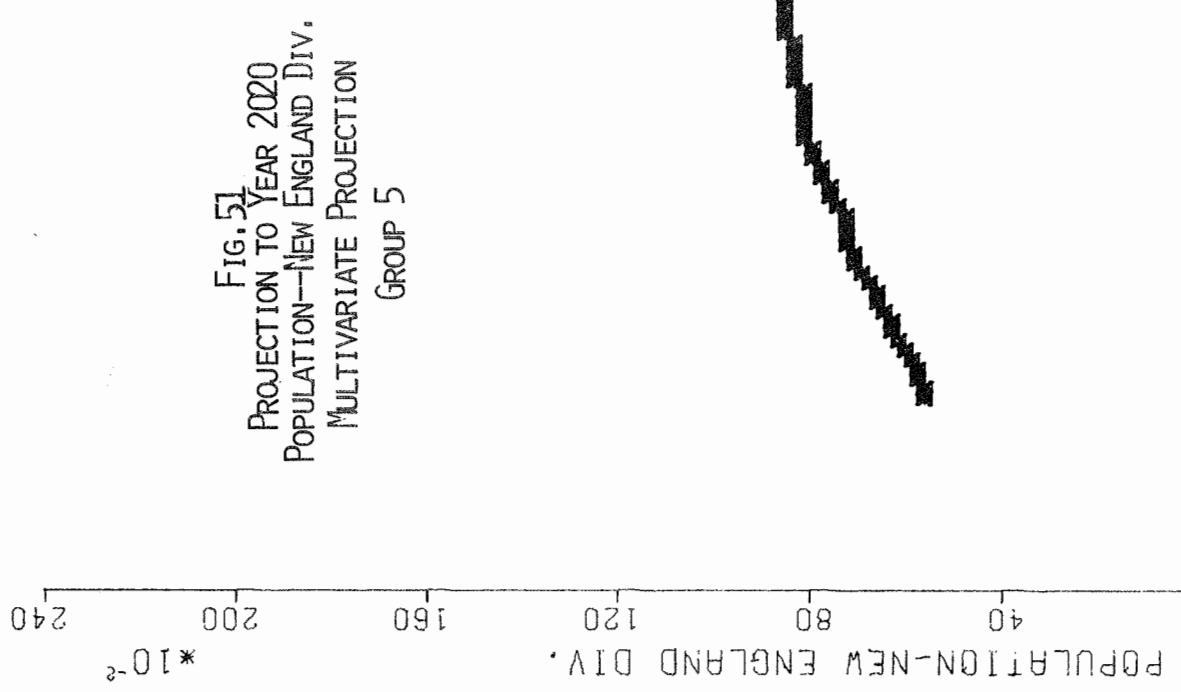


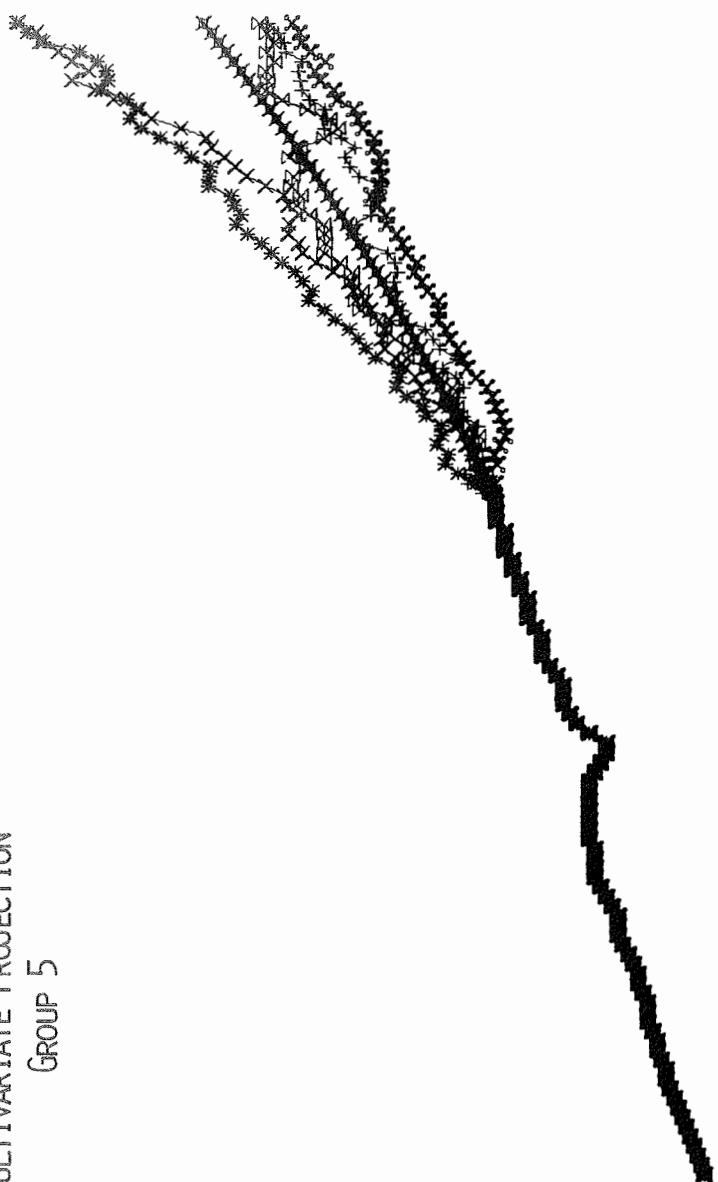
FIG. 50
PROJECTION TO YEAR 2020
COMMERCIAL PAPER RATES
MULTIVARIATE PROJECTION
GROUP 4



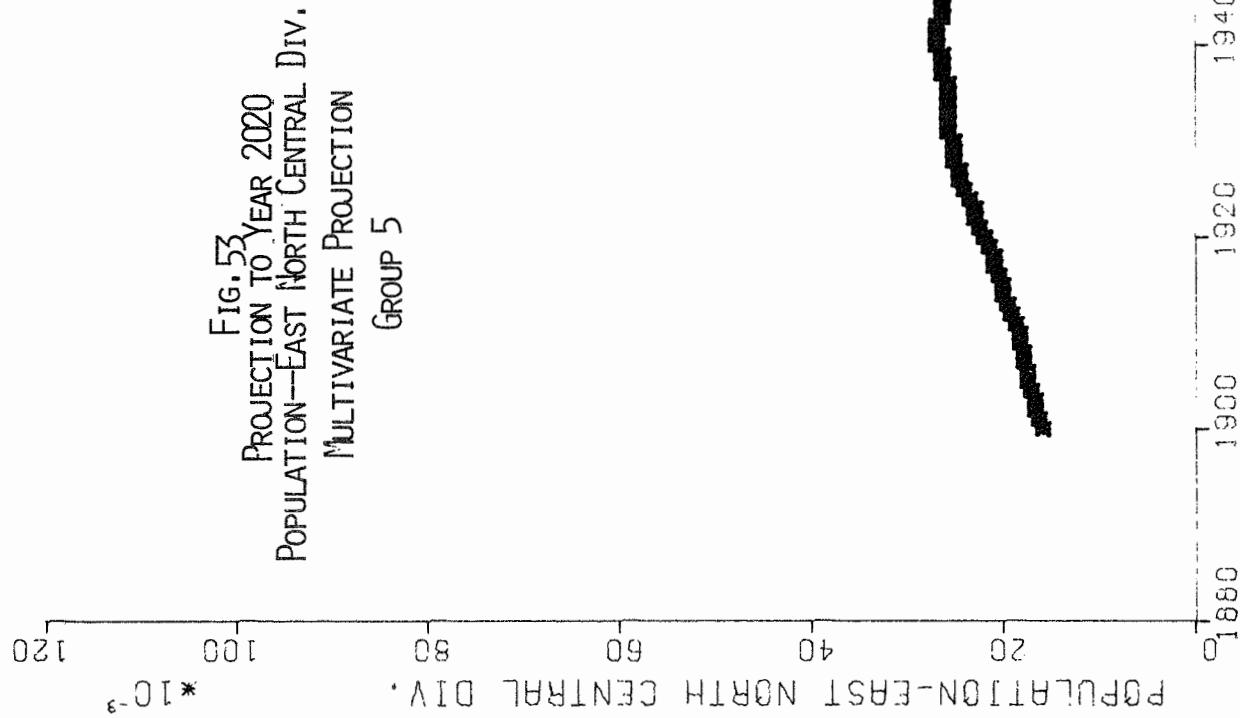


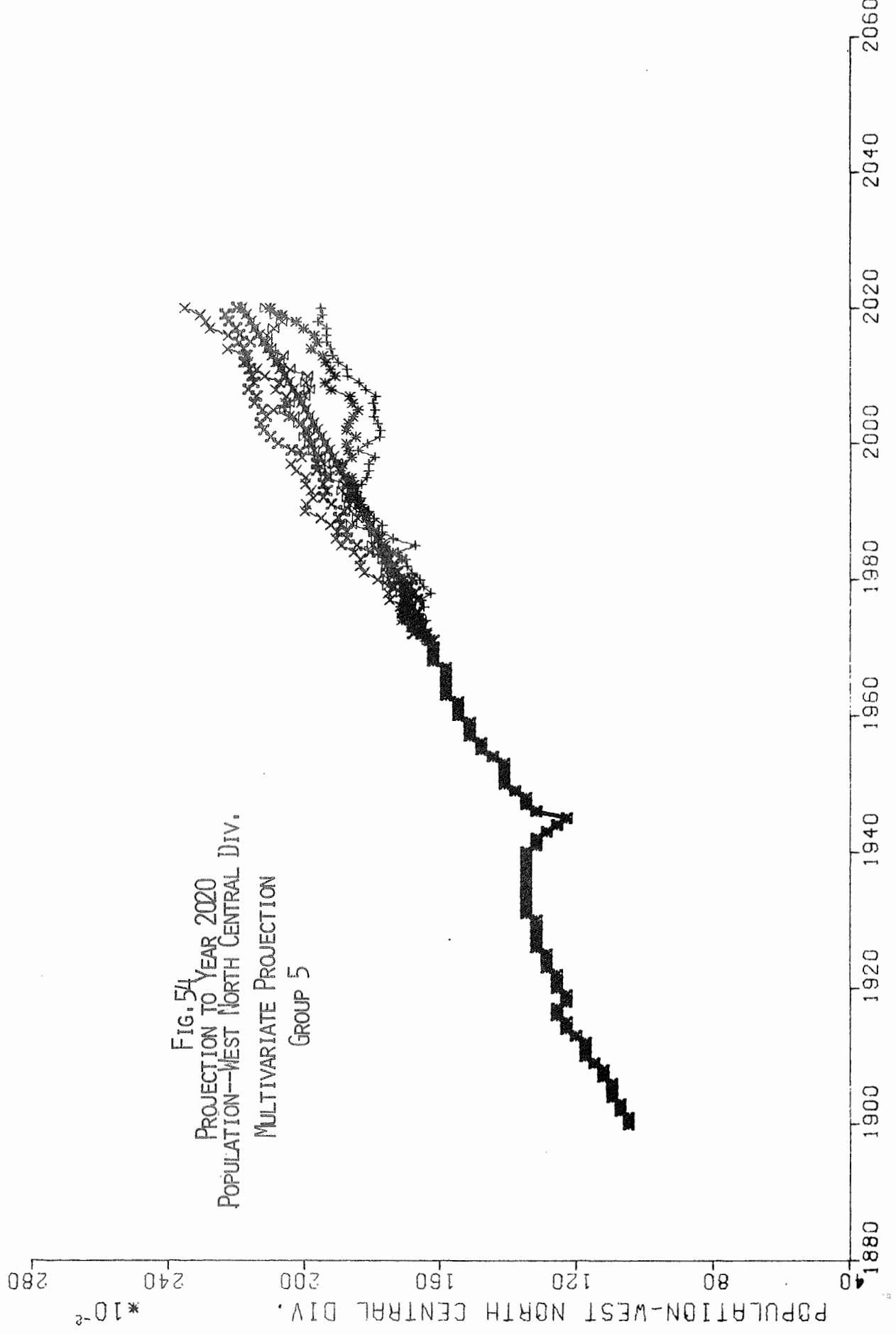
POPULATION-MIDDLE ATLANTIC DIV.
 $*10^3$

FIG. 52
PROJECTION TO YEAR 2020
POPULATION-MIDDLE ATLANTIC DIV.
MULTIVARIATE PROJECTION
GROUP 5



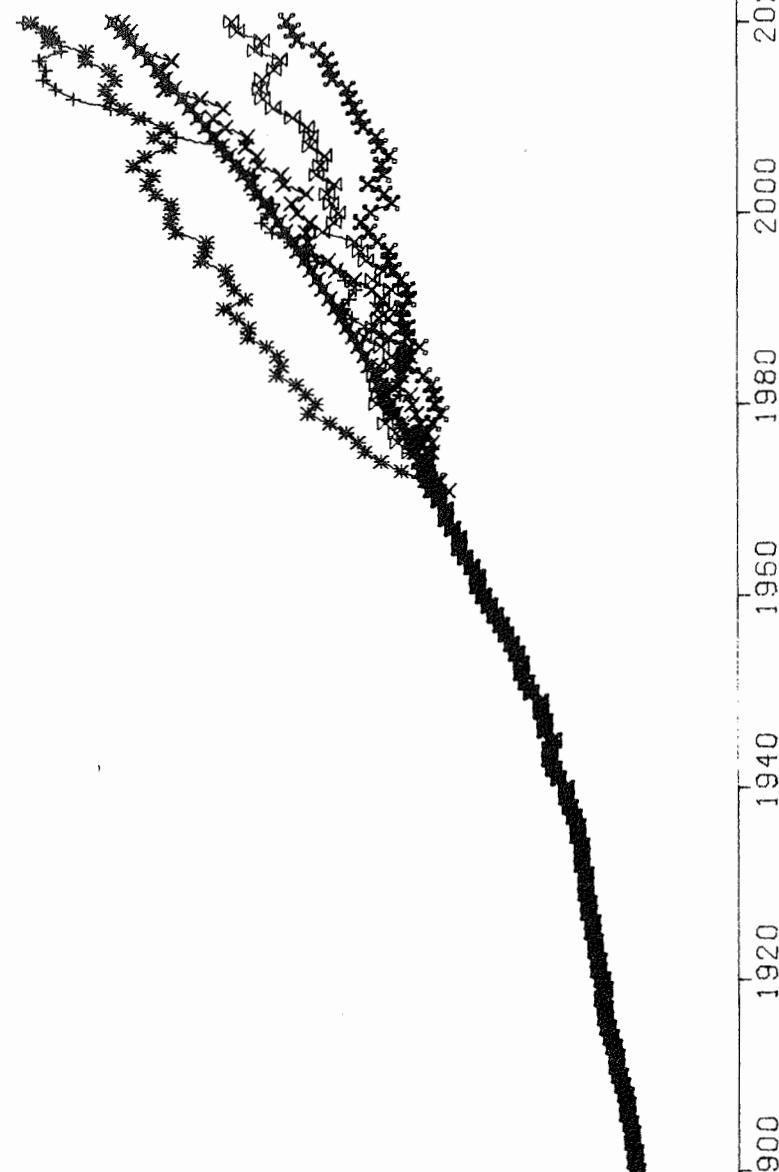
1880 1900 1920 1940 1960 1980 2000 2020 2040 2060



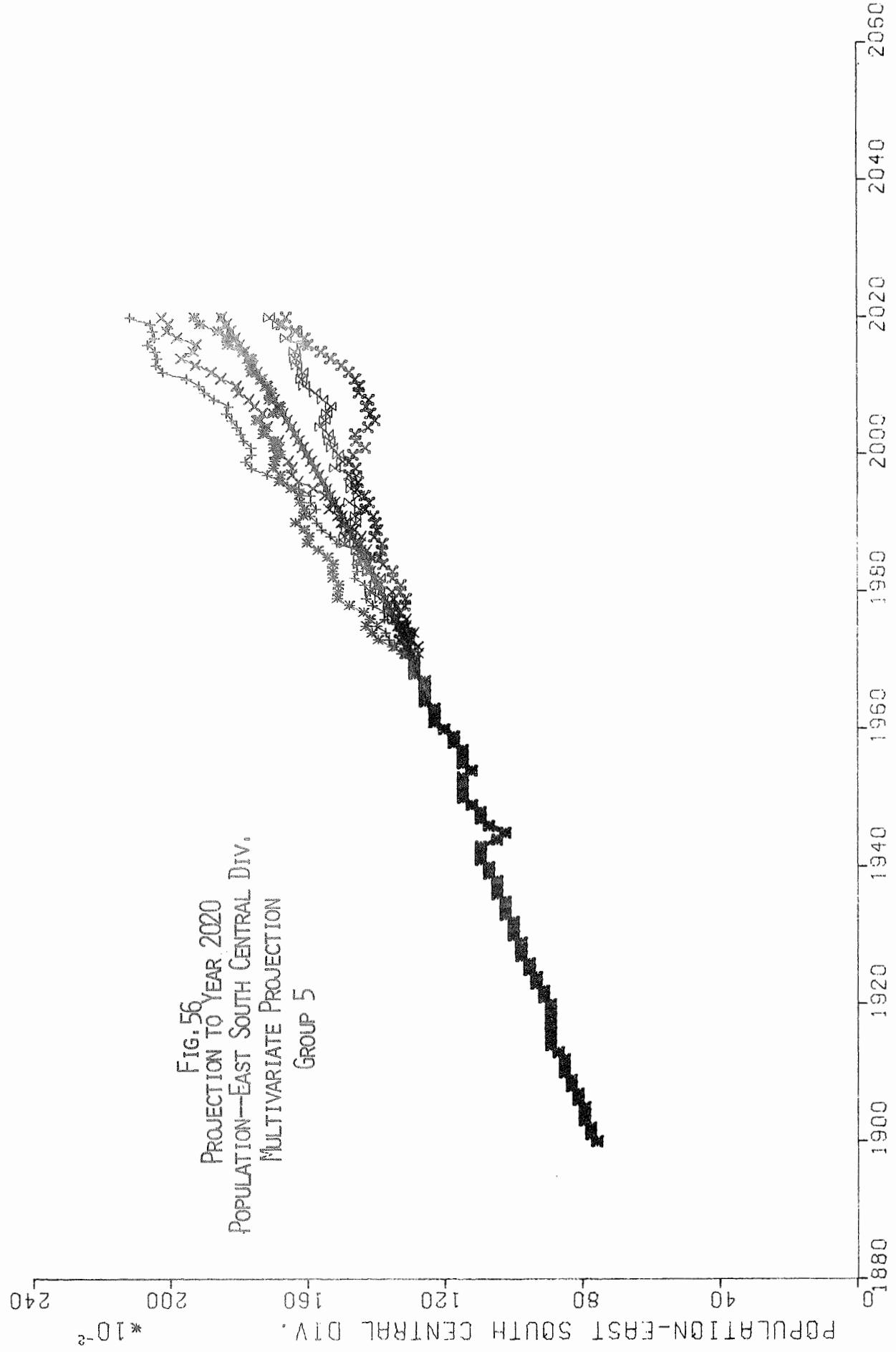


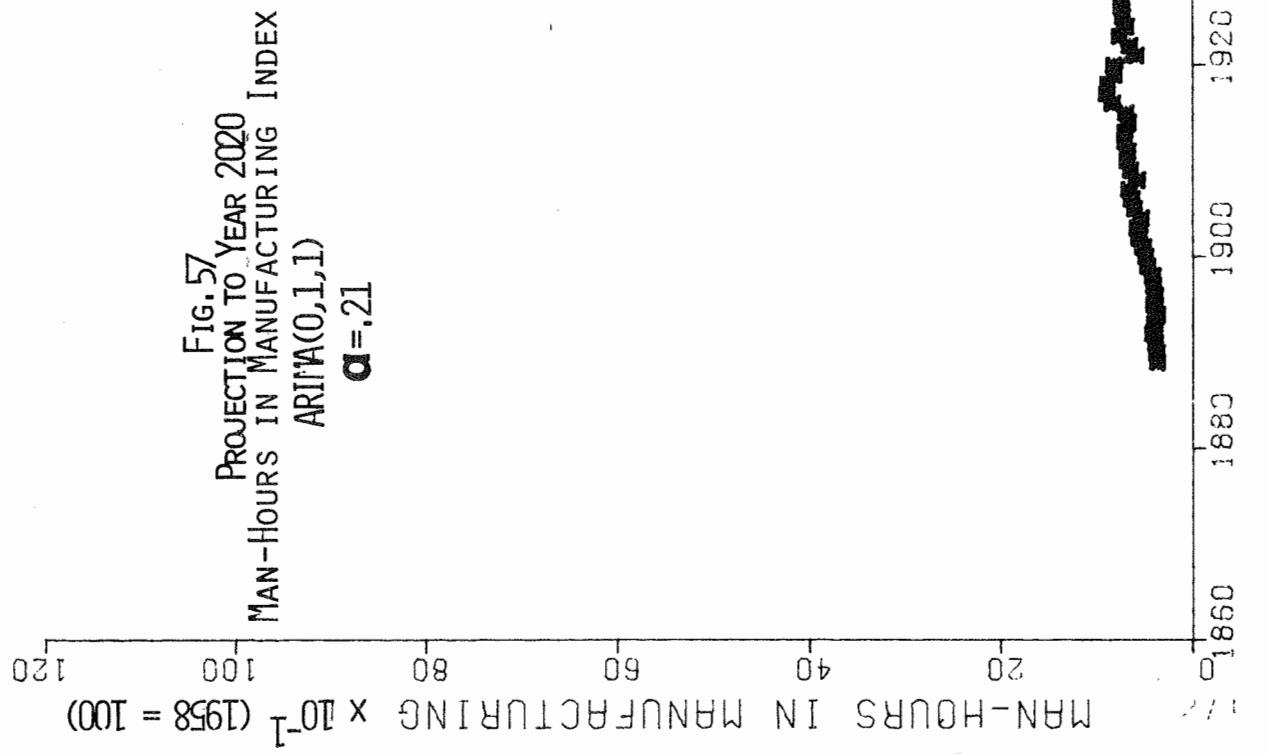
POPULATION-SOUTH ATLANTIC DIV.
 $*10^{-3}$

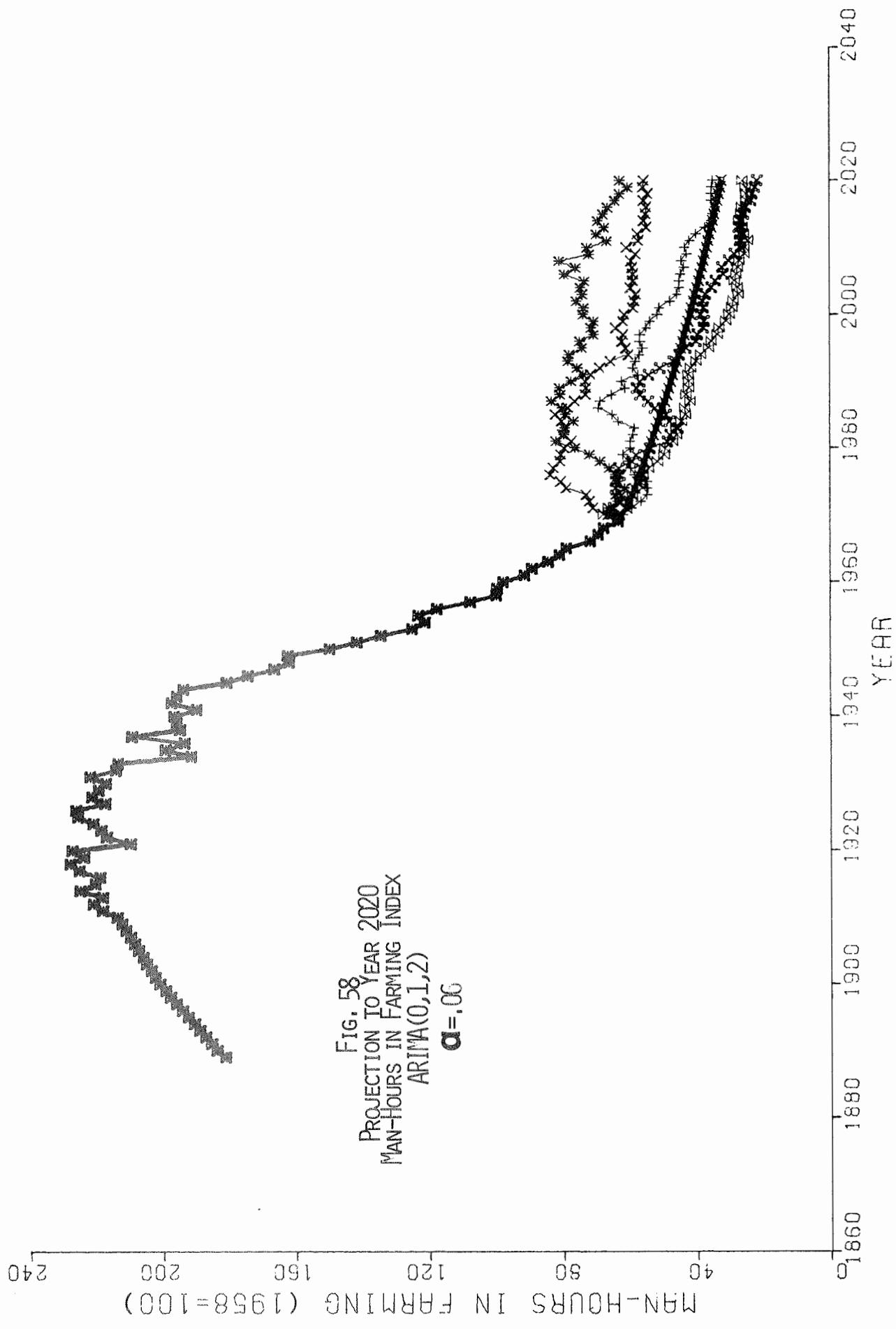
FIG. 55
PROJECTION TO YEAR 2020
POPULATION--SOUTH ATLANTIC DIV.
MULTIVARIATE PROJECTION
GROUP 5



1880 1900 1920 1940 1960 1980 2000 2020 2040 2060







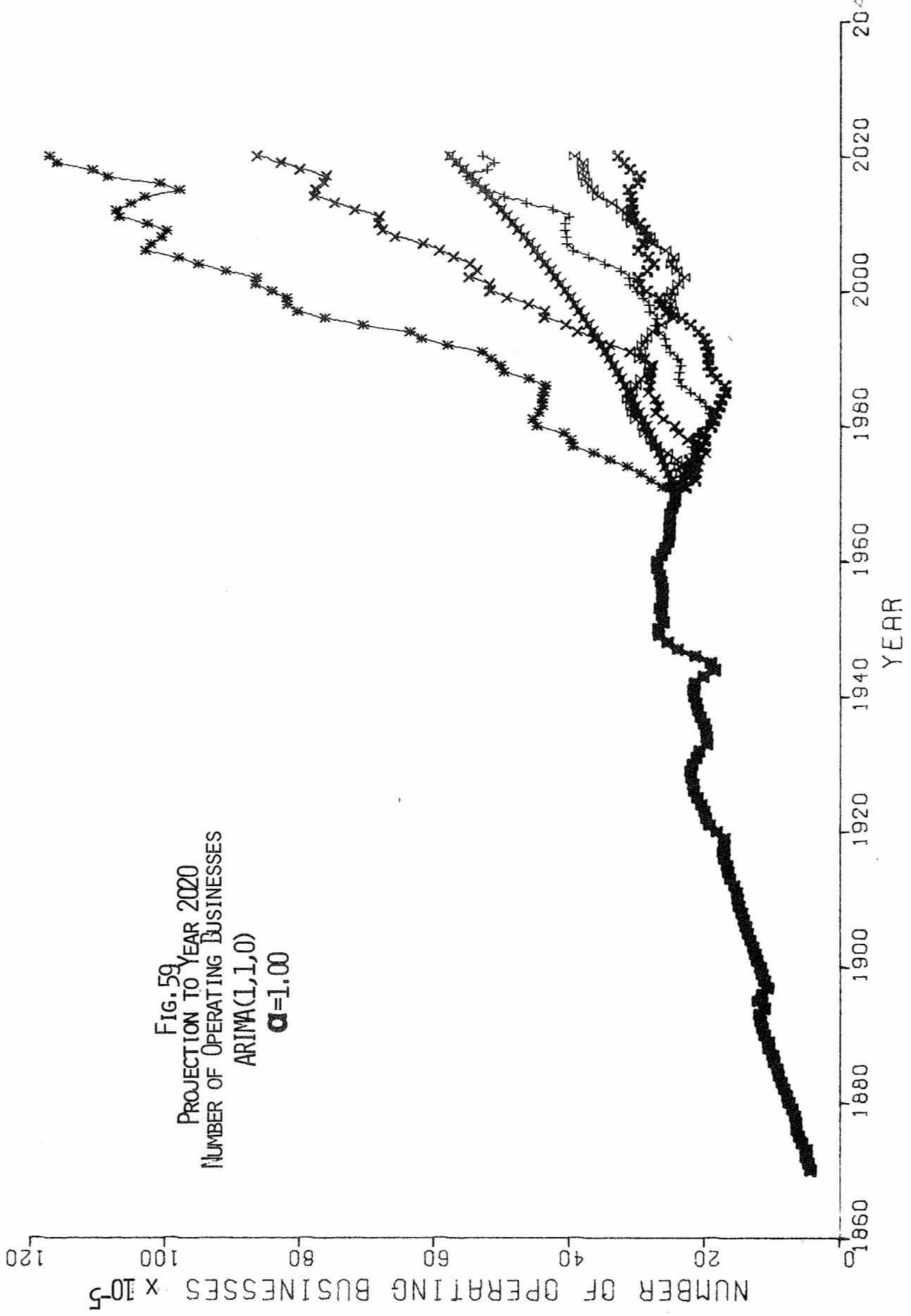
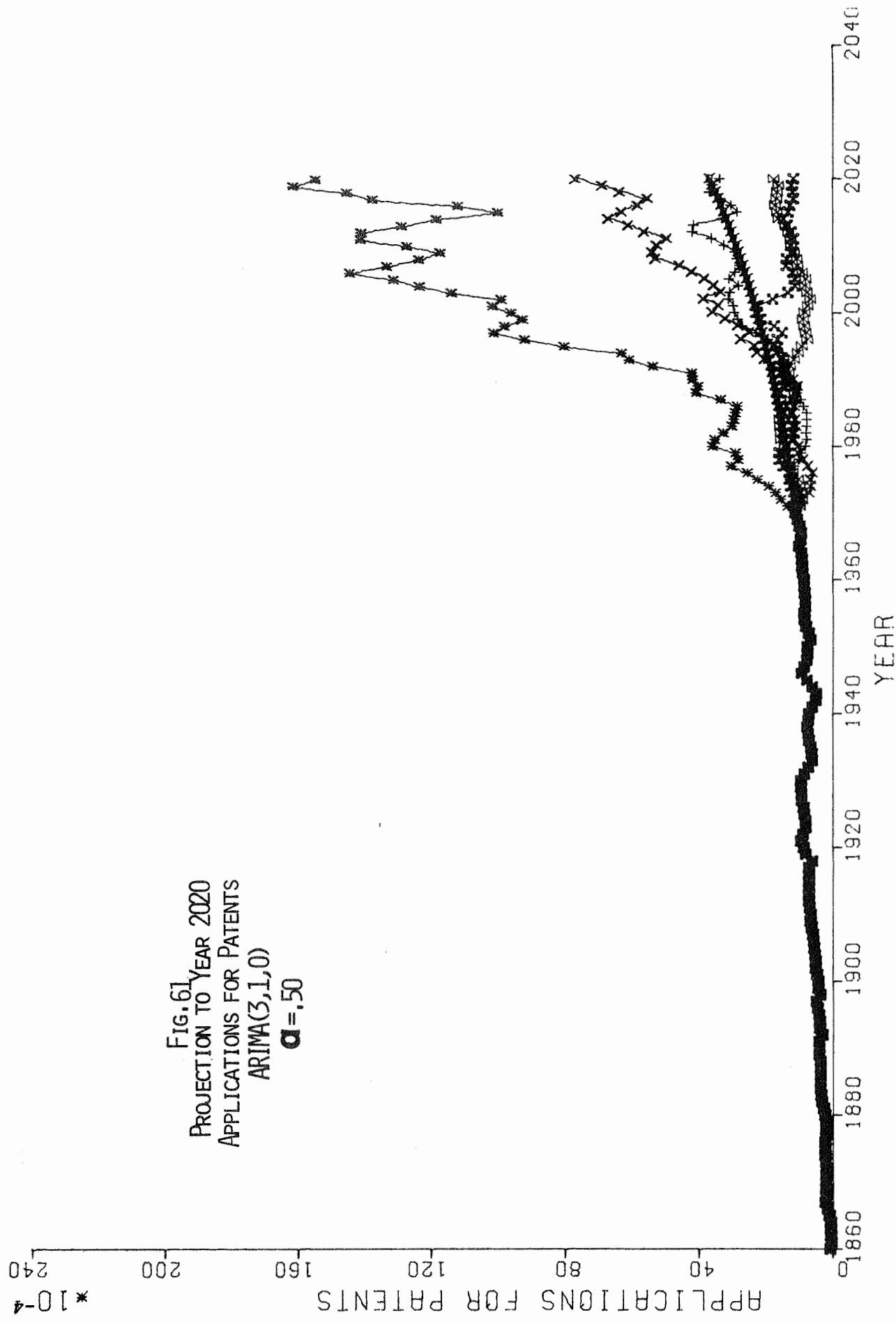
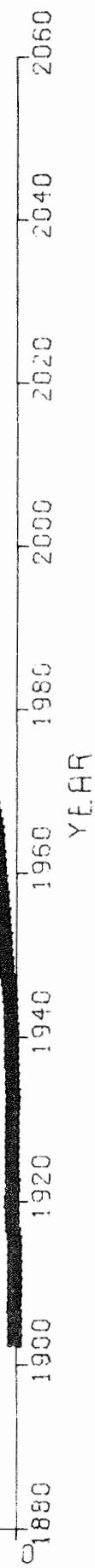
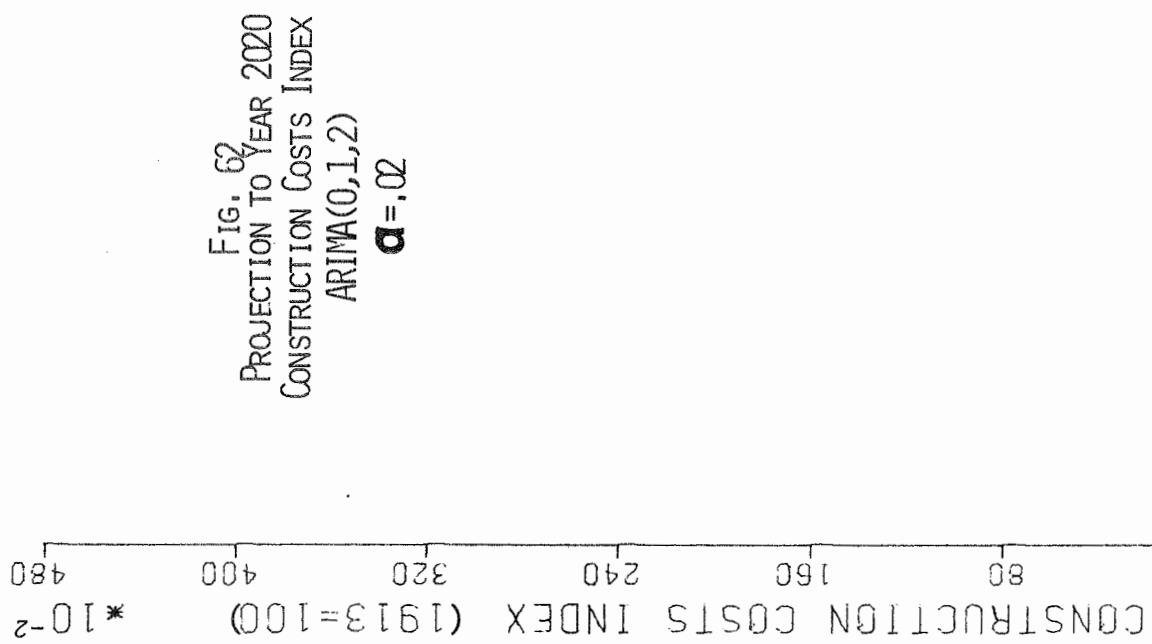




FIG. 60
PROJECTION TO YEAR 2020
UNEMPLOYMENT PERCENTAGE RATE
UPPER BOUND=100
ARIMA(0,1,2)
 $\alpha = .86$

Fig. 61
PROJECTION TO YEAR 2020
APPLICATIONS FOR PATENTS
ARIMA(3,1,0)
 $\alpha = .50$





APPENDIX A
DESCRIPTION OF COMPUTER PROGRAM SVPROJ

INTRODUCTION

1. Origin of Program. This program was developed by Shin Chang of the Center for Research in Water Resources, the University of Texas at Austin under the supervision of Leo R. Beard, Technical Director. Subroutines obtained from other sources are:

- a. FTARPS, FTMAPS, FTAUTO, FTCAST, FTGENI, MDCH and their supporting subroutines FTFUNC, GGNOR, GGUI, LEQTIF, LUDATF, LUEL MF, MERF, MERFI, MGAMMA, UERTST, VABMXF and ZSYSTM. All are included in the International Mathematics and Statistical Library (IMSL) computer package.
- b. PLOT was developed by Mr. David Ford and Mr. David Lott of the Center for Research in Water Resources, The University of Texas at Austin.

2. Capability of the SVPROJ Program. SVPROJ is capable of performing the following jobs:

- a. Identification of the degree of differencing, d, and number of parameters p and q of the best-fit ARIMA(p,d,q) model for time series of a single variable.
- b. Estimation of parameters of an identified ARIMA(p,d,q) model.
- c. Generation of 51 independent sets of projections using the best-fit ARIMA(p,d,q) model for a single variable.
- d. Plotting of up to 7 selected projections on one chart.

3. Hardware and Software Requirements. This program has been developed and tested on the CDC 64/6600 computer system. No tapes or disks are required. The Fortran IV complier and the IMSL library are used.

DESCRIPTION OF PROGRAM

4. Program Organization. The SVPROJ program consists of a main program and 23 subroutines, of which 18 are subroutines from IMSL. All input

data and input control parameters are read in the main program from punched cards. The program will make a normal exit (unless errors occur) only in the main program.

5. Method of Computation. The computational procedures and equations used in this program are described in chapter 4 of the main report under the title "Use of the ARIMA Model."

6. Description of Subroutines and Variables. The function of each subroutine and the definitions of variables in each subroutine are given in the program listings.

PROGRAM USE

7. Dimensional Limitations of the Program. For practical purpose, the following limitations are set in the program:

<u>Variable</u>	<u>Size</u>	<u>Definition</u>
LSPLIT	150	length of calibration series
LPROJ	100	length of projection
NSET	51	Sets of independent projections

These dimensions can be easily extended by increasing the dimension specification in related arrays, however, ultimate limitations are determined by the available central memory of the computer system used.

8. Control Parameters for Optional Jobs. The following parameters are used to control the execution of certain optional jobs in this program.

- IBOND: Controls the execution of lower and/or upper bound on the values of projections
- ISPLIT: Controls the execution of the split-record
- ISPY: Controls the execution of the model parameter identification for the calibration series
- LOGX: Controls the execution of natural logarithmic transformation of the calibration series.
- NPRINT: Controls the execution of the printed output of the selected projections (7 sets)

NPUNCH: Controls the execution of the punched-card output of the selected projections (7 sets).

NPLOT: Controls the execution of the plotted output of the selected projections (7 sets sets, or 5 sets).

9. Preparation Of Input Variables.

<u>Variable</u>		<u>Description</u>
IBOND:	-1	a nonzero lower bound control is used
(Requires LOGX=1)	0	a zero lower bound is used
	1	a zero lower bound and a nonzero upper bound is used
	2	both nonzero lower and upper bounds are used
ID	integer	degree of differencing, d, in an ARIMA (p,d,q) model
IP	integer	number of autoregressive parameters, p, in an ARIMA (p,d,q) model
IQ	integer	number of moving average parameters, q, in an ARIMA (p,d,q) model
ISPLIT	1	split-record test of the calibrated-ARIMA model is desired,
	0	not desired
ISPY	1	identification of ID, IP, and IQ for an ARIMA model by the program is desired
	0	ID, IP, IQ are provided by the user through input.
IYRB		beginning year of record of the calibration series
IYRE		ending year of record of the calibration series
IYRP		ending year of record of the projection series
LOGX	1	natural logarithmic transformation of the variable is desired,
NPLOT	0	no plot is desired as output
	1	plots of data and 7 projections depicted in one chart are desired
	2	plots of data and 7 projections, and data and 5 projections, depicted in 2 charts are desired

	3	plots of data and 5 projections depicted in one chart are desired
NPRINT	1	7 sets of projections of the variable being analyzed will be printed as output
	0	not desired
NPUNCH	1	7 sets of projections of the variable being analyzed will be punched as output
	0	not desired
NV	integer	number of variables to be analyzed se- quentially in a multiple job run
YLABEL (8)	character	name of the variable
Z (150)	real	a vector of N elements containing data of the calibration series, and N is computed internally as IYRE-IYRB + 1
ZLO	real	lower bound of the variable being analyzed
ZUP	real	upper bound of the variable being analyzed

10. Preparation of Input Cards. The following eight types of input cards are required.

<u>Card</u>	<u>Format</u>	<u>Variable(s)</u>
A	16I5	NV
B	8A10	YLABEL
C	16I5	IYRB, IYRE, IYRP
D	8A10	IFMT
E	IFMT	Z
F	16I5	ISPLIT, LOGX, ISPY, IBOND, NPRINT, NPUNCH, NPLOT
G	16I5	IP, ID, IQ
H	2F20.4	ZUP, ZLO

It should be noted that the G card is not required if ISPY = 1, and the H card is not required if LOGX = 0 or IBOND = 0.

11. Arrangement of Input Cards.

a. Arrangement of the input deck for a single job is:

A, B, C, D, E, E ... E, F, G, H

b. Arrangement of the input deck for a multiple job (maximum 14 jobs)

$A, B, C, D, E, \dots, E, F, G, H$, $B, C, D, E, \dots, E, F, G, H$, $\underbrace{B, C, D, E, \dots, E, F, G, H}$

12. Overview of Output. The following printed output is shown for each job:

- a. IP, ID, IQ of the identified ARIMA (p,d,q) model.
- b. α , the significance level of goodness of fit
- c. estimated parameters (θ_0 , ψ_i and θ_i) of the ARIMA (p,d,q) model

Printed and punched values of 7 sets of projections, and their plots in one curve are optional outputs.

13. Program Listing. Source program listing for SVPROJ and all its subroutines are shown in the following pages.

```

PROGRAM SVPROJ(INPUT,OUTPUT,PUNCH,TAPES=INPUT,TAPE6=OUTPUT,
  TAPET=PUNCH),
  DIMENSION Z(1,10),ALPHA(2),DARHS(15),PMASS(5),FCST(3,100),
  S1(ST(1,10),M1(M1)),YLABEL(10),IPQ(7),INDEX(5),IPM(8),
  DIMENSION ZLST(50,14),IND(14),PC(14),N(14),ND(14),IARM(5,3),
  DIMENSION G(250,9),RANGE(14),RLIMIT(14),ALFA(14),INDEX(55),
  DATA NODE/2,8,26,-44,50/
  DOUBLE PRECISION SED
C-----*
C N          NUMBER OF DATA FOR A VARIABLE
C NV         NUMBER OF VARIABLES TO BE ANALYZED
C IYRB       BEGINNING YEAR OF THE DATA
C IYRE       ENDING YEAR OF DATA
C ISPLIT     FIRST SPLITTED DATA WILL BE USED TO CALIBRATE THE MODEL
C IYSP       LENGTH OF PROJECTION
C LPROJ      SPLITTED+PROJ
C LSP        ASYMMETRIC UPPER BOUND OF Z
C ZUP        ASYMMETRIC LOWER BOUND OF Z
C ISPLIT    IF SPLIT=1,SECOND TEST IS DESIRED
C LOGX      LOGARITHMIC TRANSFORMATION OF THE ORIGINAL VARIABLE
C LOGZ      19 DESIRED
C ISPY       0, IF NOT
C           IDENTIFIED AND ID OF THE TRANSPY(D,Q) MODEL WILL BE
C IBOND     IDENTIFIED AND IQ WILL BE INPUT BY USER
C IBOND     PARAMETER TO SPECIFY THE DESIRED BOUNDS CONTROL IN Z
C           20, TO ROUND Z NOT LESS THAN ZLO
C           21, TO ROUND Z NOT LESS THAN 0
C           22, TO ROUND Z IN THE RANGE OF ZUP AND ZLO
C           23, OPTIONAL PRINTED OUTPUT OF 7 PROJECTIONS AND THE
C           MOST LIKELY FORECASTING IS DESIRED
C           24, NOT DESIRED
C           25, NO PUNCHED OUTPUT OF FORECAST AND 7 SIMULATIONS IS
C           26, PUNCHED, OUTPUT IS DESIRED
C           27, PARAMETER TO SPECIFY THE DESIRED PLOTS OF PROJECTIONS
C           28, NO PLOT IS DESIRED
C           29, PLOTS OF DATA, FORECAST, AND 9 SETS OF SIMULATIONS
C           30, CURVES FOR NPLOT(1) AND 3 ARE BOTH DESIRED
C           31, PLOTS OF DATA, FORECAST AND 5 SIMULATIONS ARE DESIRED
C           32, POSITIVE SIGNIFICANCE LEVEL IN THE SUMMARY TABLE
C           33, NEGATIVE TO GIVE ERROR INDICATION
C           34, NO SOLUTION FROM STARS
C           35, NO SOLUTION FROM STARS
C           36, NO PREDICTABLE RANGE FROM CHECK
C           37, NO PREDICTABLE RANGE FROM CHECK
C           38, NO SOLUTION FROM FTMAPS
C           39, NO ALPHAS FAIL TO PASS THE MINIMUM LEVEL
C           40, RANGE OF PROJECTION EXCEEDS RLIMT
C           41, GIVES 7 PROJECTIONS,
C           42, GIVES THE MOST LIKELY FORECASTS,
C           43, GIVES THE DATA, BUT G(J,9)=G(J,6) WHEN J IS
C           GREATER THAN N
C-----*
C NSET51    MSE=5
C           MP1=MSE+1
C           MP2=MSE+2
C           MP3=MSE+3
C           MP4=MSE+4
C           READ 1398, NV
C-----*
DO 1388 ITT=1,NV
  20  ZL=0.0
  20  ZUP=1.0E+30,ZL=-1.0E+30,YLABEL(10),JJ=1,6
  30  READ 1398, YLABEL(10)
  30  READ 1398, YYR,IYRE,IYRP
  40  NYIRE=INTH+1
  40  READ 1398, ITMT
  40  READ 1398, (Z(JJ),JJ=1,N)
  50  READ 1398, (ISPLIT,L0X,ISPY,IBOND,NPRINT,NPUNCH,NPLOT
  50  PRINT 1398, ISPLIT,L0X,ISPY,IBOND,NPRINT,NPUNCH,NPLOT
  50  IF (LISPY.FQ.0) READ 1398, IP, ID, IQ
  50  IF (IBOND.NC.0) AND (L0X.NE.0) READ 1410, ZUP,ZLO
  50  PRINT 1400, (YLABEL(JJ),JJ=1,7)
  50  DO 1010 JS1=N
  50  G(J,MP)=Z(JJ)
  50  1010 CONTINUE
  50  ZLASTL(ITT)=Z(JJ)
  50  LSPPLIT=LSPPLIT
  50  LPROJ=YRP-YYRE
  50  IF (LSPPLIT.EQ.0) GO TO 1020
  50  LSPPLIT=N+2+1
  50  LPRIJAK=LSPPLIT
  50  1020 CPUTRUE(LSPPLIT
  50  LSPPLIT=LSPPLIT
  50  C-----*
C           IF (LOGZ.NE.LOGX) GO TO 1070
C           TEHRZ=ZL
C           DO 1030 JJ=1,N
C           Z(JJ)=ZL
C           1030 CONTINUE
C           IF ((IBOND.LE.0) GO TO 1030
C           DO 1040 JJ=1,N
C           Z(JJ)=ALOG(TEMP/Z(JJ))
C           1040 CONTINUE
C           1050 CONTINUE
C           DO 1060 JJ=1,N
C           Z(JJ)=ALOG(Z(JJ))
C           1060 CONTINUE
C           1070 CONTINUE
C           SEARCH THE RANGE OF TWO CALIBRATION SERIES
C           ZMIN=Z(1)
C           ZMAX=Z(1)
C           DO 1080 JJ=2,LSPPLIT
C           ZMIN=Z(JJ)
C           ZMAX=Z(JJ)
C           1080 CONTINUE
C           RCZMAX=ZMIN
C           RCZMIN=ZMIN
C           RLIMITIT=50.0*RC
C-----*
C           KEEP GOING THE SEARCH TO GET THE RANGE OF THE WHOLE SERIES
C           AS RC AND SET THE RL FOR IDENTN ROUTINE
C           DO 1090 JJ=MSP,1
C           IF ((Z(JJ).GT.ZMAX) ZMAX=Z(JJ)
C           1090 CONTINUE
C           RCZMAX=ZMIN
C           RCZMIN=ZMIN
C           RL=50.0*RC
C-----*
C           INITAILIZE PARAMETERS
C-----*
C           IP00(1)=1
C           IP00(2)=2
C           IP00(3)=3
C           IP00(4)=4
C           IP00(6)=PROJ
  500  C-----*
  500  MSE=5
  500  MP1=MSE+1
  500  MP2=MSE+2
  500  MP3=MSE+3
  500  MP4=MSE+4
  500  READ 1398, NV

```

```

DO 1110 JPC=1,NV
  REMIND ITAPE
  DO 1220 JT=1,N
    Y(JT)=C(JT,JPC)
  1070 CONTINUE
  CALL FOCAST (Y,N,JPC)
  CC   STORE LAST MEMBERS OF SIMULATED P-C
  DO 1080 JS=1,MSET
    U(JS,JPC)=SIM(LPROJ,JS)
  1080 CONTINUE
  DO 1090 JT=1,LPROJ
    D(JT,MP1,JPC)=FC(2,JT)
  1090 CONTINUE
  CC   STORE ALL SIMULATIONS OF P-C IN DISK
  NTAP=JPC+12
  DO 1110 JS=1,MSET
    WRITE (NTAPE,1430) (SIM(I,JS),I=1,LPROJ)
  1100 CONTINUE
  1110 CONTINUE
  CC   CONSTRUCT LAST MEMBERS OF SIMULATED VARIABLE USING LAST
  MEMBERS OF SIMULATED P-C
  DO 1140 JV=1,NV
    DO 1130 JS=1,MSET
      TEMP=0.0
      DO 1120 JPC=1,NV
        TEMP=TEMP+(CJS,JPC)*Z(JV,JPC)
      1120 CONTINUE
      TEMP=TEMP/X(JV)+XMEAN(JV)
      IF (LOG(.N.EQ.0)) TEMP=10.0*TEMP
      ZL(JV,JV)=TEMP
    1130 CONTINUE
  1140 CONTINUE
  CC   COMPUTE WEIGHTED SUM OF LAST MEMBERS OF SIMULATED VARIABLE
  DO 1160 JS=1,MSET
    TEMP=0.0
    DO 1150 JV=1,NV
      TEMP=TEMP+W(JV)*ZL(JS,JV)
    1150 CONTINUE
    MTSUM(JV)=TEMP
  1160 CONTINUE
  CC   A=MTRANS( RANK MTSUM( ) FROM MAX TO MIN
  A=MTRANS(1)
  DO 1170 JS=1,MSET
    IF (MTSUM(JS).LT.0) A=MTRANS(MTSUM(JS))
  1170 CONTINUE
  A=MTRANS(10.0)
  DO 1180 JT=1,MSET
    A=MTRANS(1)
  1180 CONTINUE
  MTSUM(1)=MIN
  INDEX((1))=10
  1190 CONTINUE
  CC   CONSTRUCT SELECTED REPRESENTATIVES OF MSETS SIMULATED P-C
  DO 1230 JPC=1,NV
    NTAP=JPC+12
    REMIND NTAPE
    DO 1280 JS=1,MSET
      READ (NTAPE,1430) (SIM(I,JS),I=1,LPROJ)

```



```

C----- DIMENSION Z(1),IPDQ(4),ALPH(1),DARPS(1),PCST(1,1),
1   DIMENSION NMEAN(60),SM(18,1),MK(1),PHAS(1)
2   DIMENSION NMEAN(60)  SED
3   DOUBLE PRECISION 1E-9
4   DATA 1E-9/11/1/13/3/15/5/12B/2B/
5   SET OUTPUT PARAMETERS OF IPDQ
6
7 1010 CONTINUE
8  NIPDQ(2)
9  IP=IPDQ(3)
10 IP1=IP+1
11 IQ=MN(1)
12 IQ1=IQ+1
13 IF (IP.EQ.0.AND.IPDQ(5).EQ.0) IPDQ(5)=1
14 ID=IPDQ(5)
15 IP=MAX(IP,IQ)+1
16 MN=N
17 C
18 L=IP+1Q
19 N2=IP
20 DO 1120 I=1,N
21   WK(I)=Z(I)
22 1020 CONTINUE
23  IER=0
24 IF (IPDQ(5).EQ.0) GO TO 1050
25 1030 CONTINUE
26  N=MN-1
27  C
28  DIFFERENCE SERIES
29  DO 1040 I=1,N
30   WK(I)=WK(I+1)-WK(I)
31 1040 CONTINUE
32  IF (IPDQ(5).EQ.0.2,AND.NM1.EQ.0) N=1 GO TO 1030
33 1050 CONTINUE
34  N=MN+N
35  C
36  CALL FTIAUTO (WK,N,L,10,13,WBAR,WK,MK,MK)
37  TMP=N
38  TM=SQRT((WK(N+1))/((TMP+1)*TMP))/SORT(TMP)
39  C
40  ISET=IPDQ(7)
41  DO 1060 I=1,ISET
42   TMP=1.0
43   TEMP=TEMP*(1.0-.5) TMP=.5
44   IF (TEMP.GT.0.5) TEMP=1.0
45   TEMP=TEMP*2.0
46   TMP=-1
47   CALL MODE1 (TEMP,TMP,X,IER)
48   X=X*TP
49  NMEAN(1)=WBAR-X*STE
50 1060 CONTINUE
51  C
52  FIND DARPS,PHAS,PHAC,MNV
53  J=N+L+2
54  TEMP=1.0
55  IF (IPDQ(3).EQ.0) GO TO 1070
56  CALL FTIAPS (MK(N+1),WBAR,IPDQ(3),IPDQ(4),DARPS,PHAC,MK(J),IER)
57  TEMP=PHAC/WBAR
58  IF (IER.GT.0) GO TO 1230
59  CALL CHECK (IPDQ(3),DARPS(1),IER)
60  IF (IER.GT.0) GO TO 1240
61
62 1070 CONTINUE
63  IF (IPDQ(4).EQ.0) GO TO 1080
64  CALL FTMAPS (MK(N+1),DARPS,IPDQ(3),IPDQ(4),IPDQ(5),IPDQ(6),PHAS,MNV,MK(J),IER)
65  IF (IER.GT.0) GO TO 1250
66  IF (IPDQ(3).EQ.0) PHAC=MAR
67 1080 CONTINUE
68  CALL CHECK (IPDQ(4),PHAS(1),IER)
69  IF (IER.GT.0) GO TO 1260
70
71  C
72  IF (ID.GT.0) GO TO 1260
73  FIND ONE-STEP FORECAST ERROR
74
75  GTS 73H
76  GTS 74H
77  GTS 75H
78  GTS 76H
79  GTS 77H
80  GTS 78H
81  GTS 79H
82  GTS 80H
83  GTS 81H
84  GTS 82H
85  GTS 83H
86  GTS 84H
87  GTS 85H
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CC SEARCH HIGHEST ALFA VALUE AND STORE ITS LOCATION
C
      DO 1100 I=1,14
      ALFA(I)=ALFA(I)+(1.0*DISCON)*(NP(I)*)
1100 CONTINUE
C   ALFA(I) FROM MAX TO MIN AND STORE IN APA
      APA(1)=ALFA(1)
      DO 1120 I=1,5
      DO 1110 J=1,I
      IF (ALFA(J),LT,A) GO TO 1110
      ALFA(J)
1110 CONTINUE
      ARF1=DA
      ALFA(1)=A
      ASA
      IPARM(1,1)=NP(1)
      IPARM(1,2)=ND
      IPARM(1,3)=NO(1)
1120 CONTINUE
      RETURN
1130 FORMAT (0I4,1I1F9.3)
1140 FORMAT (4I4,1F6.2,0F16.1)
      END

SUBROUTINE PLOT (Q,N,NV,IFCR,BEGIN,NAME)
C
C   FUNCTION
C     TO PLOT UP TO 9 CURVES IN ONE CHART
C     INPUT MATRIX OF DIMENSION N BY NV
C     LENGTH OF EACH VARIABLE IN Q(I,J)
C     NUMBER OF VARIABLES IN Q(I,J) TO BE PLOTTED
C     IFCTR =SCALING FACTOR
C     YLABEL =BEGINNING DATE OF THE RECORD (USED IN X-AXIS)
C     NOTE 1. DIMENSIONS OF THE PARAMETERS IPLOT AND ISAVE (CONTROL
C           THE PRINTING SPACE ) ARE NOT CHANGEABLE
C           2. DIMENSION OF Q AND THE FIRST IF STATEMENT SHOULD BE
C              CHANGED ACCORDING TO THE SIZE OF THE MATRIX PLOTTED
C
C
      DIMENSION IPLOT(12),LIR(10),ISAVE(12),SCALE(13),Q(250,7),NAME(8)PLO 150
      DATA LBLK,IPER,1H,/1H,/1H,1H4,1H5,1H6,1M7,1MH,1H/
      DATA LTR(1H,1H,1H,1H4,1H5,1H6,1M7,1MH,1H)/
      WRITE (6,120B) (NAME(I),J=1,8)
      IF (N.GT.250 .OR. NV.GT.9) GO TO 1190
      TMP=1.0*FCR
      TMP=1.0*FCR
      WRITE (6,121B) TMP
      DO 1190 I=1,121
      ISAVE(I)=IBLANK
      IPLOT(I)=IBLANK
1190 CONTINUE
      WRITE (6,122B)
      AMIN=999999.9
      AMAX=-999999.9
      DO 1200 J=1,N
      DO 1200 I=1,N
      Q(I,J)=Q(I,J)/TMP
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      PLO 11824
      PLO 11834
      PLO 11844
      PLO 11854
      PLO 11864
      PLO 11874
      PLO 11884
      PLO 11894
      PLO 11904
      PLO 11914
      PLO 11924
      PLO 11934
      PLO 11944
      PLO 11954
      PLO 11964
      PLO 11974
      PLO 119
```

```

SUBROUTINE LEQTF (A,M,N,IA,B,DGT,WKAREA,TER)
C=LEQTF-----LIBRARY 3-----C
C  FUNCTION      = LINEAR EQUATION SOLUTION - FULL STORAGE
C  USAGE         = MODE = SPACE ECONOMYER SOLUTION,
C                  = CALL LEQTF (A,M,N,IA,B,DGT,AKRE,1R)
C  PARAMETERS    A          = INPUT MATRIX OF DIMENSION M BY N CONTAINING
C                  = THE COEFFICIENT MATRIX OF THE EQUATION
C                  AX = B.
C  ON OUTPUT, A IS REPLACED BY THE LU
C  DECOMPOSITION OF A ROWWISE PERTURBATION OF
C
C  A.
C  N          = NUMBER OF RIGHT-HAND SIDES. (INPUT)
C  M          = NUMBER OF A AND NUMBER OF ROWS IN B. (INPUT)
C  IA         = NUMBER OF ROWS IN THE DIMENSION STATEMENT. (INPUT)
C  B          = INPUT MATRIX OF DIMENSION N BY M CONTAINING
C                  = THE RIGHT-HAND SIDES OF THE EQUATION AX = B.
C  DGT        = ON OUTPUT, THE N BY M SOLUTION X REPLACES B.
C  TER         = INPUT OPTION.
C                  IF DGT IS GREATER THAN 0, THE ELEMENTS OF
C                  A AND B ARE ASSUMED TO BE CORRECT TO 10GT
C                  DECIMAL DIGITS, AND THE ROUTINE PERFORMS
C                  AN ACCURACY TEST.
C                  IF DGT EQUALS ZERO, THE ACCURACY TEST IS
C                  BYPASSED.
C  WKAREA     = WORK AREA OF DIMENSION GREATER THAN OR EQUAL
C                  TO N.
C  IER        = ERROR PARAMETER
C                  TERMINAL ERROR = 1200N.
C                  N = 1 INDICATES THAT A IS ALGORITHMICALLY
C                  SINGULAR. SEE THE CHAPTER L PRELUDE).
C  WARNING     = 12N.
C  N = 2 INDICATES THAT THE ACCURACY TEST
C  FAILED.
C  THE COMPUTED SOLUTION MAY BE IN FRROR BY
C  MORE THAN CAN BE ACCOUNTED FOR BY
C  THE UNCERTAINTY OF THE DATA.
C  THIS WARNING CAN BE PRODUCED ONLY IF
C  DGT IS GREATER THAN 0 ON INPUT.
C  SEE CHAPTER L PRELUDE FOR FURTHER
C  DISCUSSION.
C
C  PRECISION   = SINGLE
C  READ, TBL ROUTINES = LUDATP, LUDEFNP, UERTST
C  LANGUAGE    = FORTRAN
C
C  LATEST REVISION = AUGUST 15, 1973
C  DIMENSION   AT(IA,1),BK(IA,1),WKAREA(1)
C

```


9885 RETURN
END

MFR1280
MFR1290

DATA
1

P1(1)/2.431752442E+02/,
P1(2)*2.617215836561E+02/,
P1(3)**9.2226137226015E+02/,
P1(4)*5.17653098235E+02/,
P1(5)*-7.7418600711322E+01/,
P1(6)/-2.2885307161E+01/,
P1(7)/-4.128531586077E+01/,
P1(8)/8.568928613E+01/,
Q1(1)/3.778372484639E+02/,
Q1(2)/-9.51325976772E+02/,
Q1(3)/6.4607536202087E+02/,
Q1(4)/2.235834902694E+02/,
Q1(5)/-1.455196658635E+02/,
Q1(6)/4.85772921616E+01/,
Q1(7)/4.564611751594E+01/,
COEFFICIENTS FOR MINIMAX,
APPROXIMATION TO LN(GAMMA(X)),
1.5 .LE. X .LE. q,0

P2(1)/5.2680325594E+03/,
P2(2)/1.5553695540030E+03/,
P2(3)/-1.203173008871E+04/,
P2(4)/-8.64629426352E+04/,
P2(5)/-1.50850267467E+05/,
P2(6)/-1.513631834157E+05/,
P2(7)/5.1559376176088E+05/,
Q2(1)/6.4607536202087E+02/,
Q2(2)/-2.235834902694E+02/,
Q2(3)/-1.455196658635E+02/,
Q2(4)/2.22966244056E+02/,
Q2(5)/-1.20555386254E+02/,
Q2(6)/5.2622663158011E+01/,
Q2(7)/6.9832741405755E+01/,
Q2(8)/-1.589321490129E+02/,
COEFFICIENTS FOR MINIMAX,
APPROXIMATION TO LN(GAMMA(X)),
4.0 .LE. X .LE. 12.0

P3(1)/-1.9713015669E+01/,
P3(2)/-8.31075135385E+01/,
P3(3)/1.91935153425983E+01/,
P3(4)/-4.81607192773.6E+01/,
P3(5)/-2.44312736932E+01/,
P3(6)/-2.407946880173E+01/,
P3(7)/-1.037701651732E+01/,
P3(8)/-9.4270281420E+05/,
P3(9)/1.91935153425983E+01/,
P3(10)/-1.146827473.6E+01/,
P3(11)/-1.119354116233E+01/,
P3(12)/-4.04459282147E+01/,
P3(13)/-4.353197488437E+01/,
P3(14)/-2.902611111111111E+01/,
P3(15)/-2.91577595880E+01/,
COEFFICIENTS FOR MINIMAX,
APPROXIMATION TO LN(GAMMA(X)),
12.0 .LE. X .LE. 14.0

P4(1)/9.189353328467E+01/,
P4(2)/6.3333333333367E+01/,
P4(3)/-2.7777776515151E+01/,
P4(4)/-2.958649764E+01/,
P4(5)/-5.822999631684E+01/,
XNP/37777777777777777778/,
P1/3.1415926535869/

DATA
1

MGM1010
MGM1020
MGM1030
MGM1040
MGM1050
MGM1060
MGM1070
MGM1080
MGM1090
MGM1100
MGM1110
MGM1120
MGM1130
MGM1140
MGM1150
MGM1160
MGM1170
MGM1180

DATA
1

MGM1200
MGM1210
MGM1220
MGM1230
MGM1240
MGM1250
MGM1260
MGM1270
MGM1280
MGM1290
MGM1300
MGM1310
MGM1320
MGM1330
MGM1340
MGM1350
MGM1360
MGM1370
MGM1380
MGM1390
MGM1400
MGM1410
MGM1420
MGM1430
MGM1440
MGM1450
MGM1460
MGM1470
MGM1480
MGM1490
MGM1500
MGM1510
MGM1520
MGM1530
MGM1540
MGM1550
MGM1560
MGM1570
MGM1580
MGM1590
MGM1600
MGM1610
MGM1620
MGM1630
MGM1640
MGM1650
MGM1660
MGM1670
MGM1680


```

IF(IQUIT .EQ. 0) GO TO 140
1   I = K
2   I=IPART(I)
3   ITEMP = W(LKSUB+1) + PH1
4   HODD = X(ITEMP)
5   ETAB=0.01ABSH(MOLD)
6   IF (ABS(HODD).LT.PREC) ETABDELTAA
7   H=AMIN1(FMAX,FJ)
8   IFHLLT PREC) HODPREC
9   X(CITMP,J)HODH
10  IF (K .LE. 1) GO TO 45
11  KK = 2
12  GO TO 15
13  PPLUSF(X,K,PAR)
14  TDPNPUSSE
15  W(AKP1)HODD
16  X((ITMP-1)*HODD
17  IF (I .LT. 1) GO TO 46
18  IF (K .LT. N) GO TO 66
19  IP=PART+N
20  IF (ABS(WAITP)) .EQ. ZERO) GO TO 80
21  IF (ABS(WAITP)) .EQ. WAITP + X(ITEMP) + X(ITEMP)
22  IF (E = WAITP) + WAITP) / WAITP) * MAC(PK)
23  W(KK)=W(KN)+W(APK)*X(JSUB)
24  CONTINUE
25  K=LKSUB+K
26  LOOK(WAKL) + PH1
27  KWALKONK
28  IP=PART+N
29  DERNAX ABS(X(WAITP))
30  KPLUS = H+1
31  DO 65 I = KPLUS,N
32  TEST ABS(WAITP+T)
33  IF(TEST .LE. DEMAX) GO TO 65
34  DEMAX = TEST
35  KMAX=1
36  CONTINUE
37  IF(LOCK .EQ. KMAX) GO TO 75
38  LMAX=LKSUB+KMAX
39  WAKL=LMAX(LMAX)
40  WALKMAXLOOK
41  IP=PART+KMAX
42  XTEMP= W(AIP)
43  IP=PART+K
44  W(AIP)=H(IPK)
45  W(AIP)=TEMP
46  IF(K .LT. 2) GO TO 75
47  KMNMAX=1
48  I1 = 0
49  DO 76 IM1,KMN
50  L=(I1)*(N2+1)/2+1
51  J=I1
52  CONTINUE
53  IF ( ABS(WAITP) ) .NE. ZERO) GO TO 90
54
55  IF (ERROR .EQ. 1) GO TO 135
56  DO 85 I=1,N
57  ZSYTH C
58  ZSYTH C
59  ZSYTH C
60  ZSYTH C
61  ZSYTH C
62  ZSYTH C
63  ZSYTH C
64  ZSYTH C
65  ZSYTH C
66  ZSYTH C
67  ZSYTH C
68  ZSYTH C
69  ZSYTH C
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174 ZSYTH C
175 ZSYTH C

```

```

145 DO 150 I = 1,N
      IP1=PAR(1,IP)
      HATIP=PAR(1,IP)
      PMAX=MAX(PMAX, ABS(WA(IP)))
146 CONTINUE
C 147 K=1
      DO 166 I=1,N
          W(I)=X(I)
          FP=(ABS(X(I))) .GT. PR) GO TO 160
          K=2
          W(I)=ZERO
166 CONTINUE
      IP(K)=2, 1) GO TO 193
      KK = 1
      GO TO 205
168 IF(FMAX .LT. TEST) GO TO 198
      NOTE THAT SMALL COMPONENTS ARE SET
      TO ZERO ONLY IF THE NORM OF THE
      FUNCTION VECTOR IS REDUCED AS A
      RESULT OF THIS PROCESS.
C
      DO 170 I=1,N
          X(I)=WA(I)
170 CONTINUE
      IF (N .GT. 1) GO TO 175
      HATIP=PAR(2) = WA(TMP+2)
      GO TO 185
175 DO 186 I = 1,N
      HATIP=PAR(I) = WA(TMP+I)
186 CONTINUE
185 PMAXTEST
C 190 K=1
      DO 195 ITEST=0
195 ITEST=0
      DO 206 I=1,N
          W(I)=X(I)
          IF ((ABS(X(I)) .LE. PR) GO TO 206
          LX(I)+P1
          JX(I)+P1
          IP(L,P0,J) GO TO 208
          HATIP=MSIGN(I,J)*MAXB(IAB8(L),IA88(J))
196 CONTINUE
      KK = 2
      GO TO 226
200 CONTINUE
      IP(K)=ED, 1) GO TO 235
      KK = 2
205 TESTZERO,
      IF (N .GT. 1) GO TO 210
      HATIP=PAR(1,IP)
      TEST=MAX(1,TEST,ABS(WA(IP+2)))
      GO TO 226
210 DO 215 I=1,N
          IA1IMP(I,
          HATIP=PAR(1,IP),
          TEST=MAX(1,TEST,ABS(WA(IP+2)))
215 CONTINUE
216 GO TO 165,225, KK
225 IF(FMAX .LT. TEST) GO TO 235
      NOTE THAT MEAN=INTEGER COMPONENTS
      ARE SET TO BE INTEGERS ONLY IF THE
      NORM OF THE FUNCTION VECTOR IS THE
      REDUCED AS A RESULT OF THIS PROCESS.

```

APPENDIX B
DESCRIPTION OF COMPUTER PROGRAM MVPROJ

INTRODUCTION

1. Origin of Program. This program was developed by Shin Chang of The Center for Research in Water Resources, The University of Texas at Austin under the supervision of Leo R. Beard, Technical Director. Subroutines obtained from other sources are BMDOIM of Biomedical Computer Programs (BMD) and all those subroutines used in Program SVPProj.

2. Capability of the MVPROJ Program. MVPROJ is capable of performing the following jobs:

- a. Construction of the principal components of a set of original variables.
- b. For each of these principal components:
 1. Identification of p, d, and q of the best-fit ARIMA(p,d,q) model for each principal component being analyzed.
 2. Estimation of parameters for the models identified in 1.
 3. Generation of 51 independent sets of projections for each principal component.
- c. Computation of 51 projections for each of the original variables as functions of projections of all principal components.
- d. Plotting up to 7 selected projections of each of the original variables in one chart for each variable.

3. Hardware and Software Requirements. This program has been developed and tested on the CDC 64/6600 computer system. Thirteen disk files are required. The Fortran IV compiler and the IMSL library are used.

DESCRIPTION OF THE PROGRAM

4. Program Organization. The MVPROJ program consists of a main program and 25 subroutines. All input data and input control parameters are read

in the main program from punched cards. The program will make a normal exit (unless errors occur) only in the main program.

5. Method of Computation. The computation procedures and equations used in this program are described in Chapter 5 of the main report under the title "Multivariate Projection Model."

6. Description of Subroutines and Variables. The function of each subroutine and the definitions of variables in each subroutine are given in the program listings. However, subroutines used in SVPProj are not repeated here.

7. Dimensional Limitations of the Program. For practical purpose, the following limitations are set in the program:

<u>Variable</u>	<u>Size</u>	<u>Definition</u>
NV	12	number of variables
LSPLIT	150	length of calibration variables
LPROJ	100	length of projection
NSET	51	sets of independent projections

These dimensions can be easily extended by increasing the dimension specification in related arrays, however, ultimate limitations are determined by the available control memory of the computer system used.

8. Control Parameters for Optional Jobs. The following parameters are used to control the execution of certain optional jobs in the program.

ISPLIT = controls the execution of the split-record test
LOGX = Controls the execution of the logarithmic (base 10)
 transformation of the original variables
NPRINT = Controls the execution of the printed output of the
 selected projections (7 sets)
NPUNCH = Controls the execution of the punched-card output of the
 selected projections (7 sets)
NPILOT = Controls the execution of the plotted output of the
 selected projections (7 sets or 5 sets)

9. Preparation of Input Variables.

<u>Variable</u>		<u>Description</u>
IFMT	character	variable format for input data x
ISPLIT	1	split-record test is desired,
	0	not desired
IYRB	integer	beginning year of record of the calibration series
IYRE	integer	ending year of record of the calibration series
IYRP	integer	ending year of projection
LOGX	1	logarithmic (base 10) transformation of the variables is desired
	2	not desired
NPLOT	0	no plot is desired as output
	1	plots of data and 7 projections depicted in one chart are desired
	2	plots of data and 7 projections, and data and 5 projections, depicted in 2 charts are desired
	3	plots of data and 5 projections depicted in one chart are desired
NV	integer	number of variables
WT (12)	real	relative weights to be applied to the last member of projection of each variable for ranking the 51 sets of projections.
YLABEL (8)	character	name of variable to be appeared in the chart
X (150, 12)	real	input data of calibration series, the first dimension refers to time and the second dimension refers to variables, (X(j,i) = jth record of ith variable)

10. Preparation of Input Cards. The following types of input cards are required.

<u>Card</u>	<u>Format</u>	<u>Variable(s)</u>
A	16I5	NV, IYRB, IYRE, IYRP, LOGX, ISPLIT, NPRINT, NPUNCH, NPLOT

B	12F5.0	WT
C	8A10	IFMT
D	IFMT	X
E	—	job card for BMD01M
F	—	label card for BMD01M
G	IFMT	data format card for BMD01M
H	8A10	YLABEL

Preparation of E and F cards of BMD01M:

E card (Job card):

col. 1-6	Problem (mandatory)
col. 7-12	Alphanumeric job code
col. 13,14	Number of original variables ($2 \leq NV \leq 12$)
col. 15-17	Length of calibration variables ($3 \leq LSPLIT \leq 150$)
col. 18-20	Blank
col. 21-23	Yes
col. 24-28	Blank
col. 29-30	Number of variables to be labeled
col. 31-68	Blank
col. 68-70	07 (data input is from logical BCD Tape 7)
col. 71,72	K Number of Variable Format Cards ($1 \leq K \leq 10$)

F Card (Labels card):

col. 1-6	LABELS (mandatory)
col. 7-10	The number of the variable to be named. The number must be right-justified
col. 11-16	The corresponding alphanumeric name
col. 17-20	The number of another variable
col. 21-26	The corresponding alphanumeric name.
⋮	⋮

col. 67-70 The number of another variable
col. 71-76 The corresponding alphanumeric name
 (up to 7 per card)

Example = The following Labels Card (F Card) is prepared for the case of Labelling $X(j,2)$ as name 01 and $X(j,5)$ as name 02.

Labels0002NAME010005NAME02.

Note 1. The number of a variable is i if the variable to be labeled is the i th variable of the input vector $X(j,i)$.

2. More than one labels card can be used.
 3. The number of labels appearing on all the labels cards must equal to the number of labels specified on the job card.

11. Arrangement of Input-deck. The following sequence of the input deck is used:

12. Overview of Output. The following printout will be shown:

- a. Normalized and transformed ($\text{LOGX} = 1$) input data
- b. Correlation coefficient matrix
- c. Eigenvalues
- d. Eigenvectors
- e. For each principal component, a summary table showing the best-fit ARIMA (p,d,q) model, estimations of parameters, level of significance.

Optional printed and/or punched output of the 7 sets of projections and chart (s) of projections will be shown upon the specification of related output control parameters.

13. Program Listing. Source program listing for MVPROJ and its subroutines are shown in the following pages. Note that subroutines shared by both programs SVPROJ and MVPROJ will be shown in appendix A only.


```

11869
11870 J=INDEX(L)
11871 DO 1190 II=1,LPROJ
11872 IF (SM(II,J)*GT.2MAX) ZMAX=SIM(II,J)
11873 IF (SM(II,J)*LT.2MIN) ZMIN=SIM(II,J)
11874 CONTINUE
11875 RPZMAX=MIN
11876 IF (RP.GT.RANGER(ITT)) RANGER(ITT)=RP
11877 IF (RP.LT.RANGER(ITT)) RANGER(ITT)=RP
11878 CC
11879 IF (RANGE(ITT).LE.RLIMIT(ITT)) GO TO 1208
11880 ALFA(ITT)=6.*6.*6
11881 CONTINUE
11882 C
11883 IDENTIFIED BY IDENTN ROUTINE USING WHOLE RECORD OF SERIZ
11884 IND(1)=N
11885 IND(2)=5*N
11886 IND(3)=5
11887 CALL IDENTN (Z,IND,RL,8,IS,ARF,IPARM,1,r1)
11888 IPARM(1)=IPARM(JTT,1)
11889 IPARM(2)=IPARM(JTT,3)
11890 IPARM(3)=IPARM(JTT,5)
11891 IF (IPARM(1).GT.0) IDPG(5)=IPARM(JTT,2)
11892 SET=0.521067054D0
11893 ALPHA(1)=0.01
11894 ALPHA(2)=0.05
11895 IS=100
11896 DO 1148 Is=1,5
11897 DAPSTIN(1)=0.0
11898 PHAS(1)=0.0
11899 CONTINUE
11900 PHACM=0.99
11901 CALL GSIMP (Z,IPDQ,SEED,ALPHA,DAPRS,PHAS,PHAC,MV,PCST,BIM,IS,WK,
11902 1,1,ER)
11903 NQ(1)=IPDQ(1)
11904 NQ(2)=IPDQ(2)
11905 NQ(3)=IPDQ(3)
11906 HD(1)=IPDQ(4)
11907 HD(2)=IPDQ(5)
11908 C
11909 RANK THE LAST MEMBERS OF SIMULATION SIM(I,J) INTO ZLAST
11910 FROM MAX TO MIN
11911 C
11912 DO 1158 J=1,1SET
11913 WK(J)=WKLPROJ(J)
11914 C
11915 SEARCH FOR THE MINIMUM AND THEN SET MIN FOR RANKING OPERAT
11916 C
11917 AMK(1)
11918 DO 1168 J=1,1SET
11919 IF ((WK(J)).LT.A) AMK(J)
11920 AMK(1)
11921 L=J
11922 C
11923 DO 1188 L=1,1SET
11924 AMIN=0.0
11925 DO 1170 J=1,1SET
11926 IF ((WK(J)).LE.A) GO TO 1178
11927 AMK(J)
11928 L=J
11929 C
11930 CONTINUE
11931 WK(1)=AMIN
11932 INDEX(L)=1
11933 C
11934 RANGE(1)=8.*8
11935 ZMAX=3*M(1,J)
11936 2*M(2,M(1,J))
11937 DO 1208 J=1,5
11938 L=J
11939 C
11940 INDEX(L)=1
11941 COMPUTE THE RANGE OF PROJECT AS RANGE
11942 IF (ISPLIT=0, SET G(J,MP4)=G(J,MP3) FOR J GREATER THAN N
11943 2438
11944 IF (ISPLIT=1, SET G(J,MP4)=G(J,MP3) FOR J LESS THAN N
11945 2448
11946 DO 1208 J=1,1LSP
11947 C
11948 G(J,MP4)=G(J,MP3)
11949 1208 CONTINUE
11950 2448
11951 2449
11952 2450
11953 2451
11954 2455
11955 2456
11956 2457
11957 2458
11958 2459
11959 2460
11960 2461
11961 2462
11962 2463
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C      I=1, IPDQ(1) IS REQUIRED FOR TIME SERIES Z.
C      I=2, INPUT LENGTH OF TIME SERIES Z.
C      I=3, NUMBER OF AUTOREGRESSIVE PARAMETERS IN THE
C      MODEL.
C      I=4, NUMBER OF MOVING AVERAGE PARAMETERS IN THE MODEL.
C      I=5, DEGREE OF DIFFERENCING OPERATION REQUIRED TO
C      OBTAIN A STATIONARY TIME SERIES.
C      I=6, FORECASTING CONTROL PARAMETER.
C      FORECASTS UP TO IPDQ(6) STEPS IN ADVANCE ARE
C      CALCULATED. IPDQ(6) MUST BE POSITIVE.
C      I=7, SIMULATION UTILITY.
C      IPDQ(7) LESS THAN OR EQUAL TO ZERO IMPLIES
C      SIMULATION NOT DESIRED.
C      IPDQ(7) POSITIVE IMPLIES IPDQ(7) SIMULATIONS OF
C      THE TIME SERIES UP TO IPDQ(6) STEPS IN ADVANCE
C      ARE DESIRED.
C      INPUT DOUBLE PRECISION NUMBER. SEED IS USED TO
C      GENERATE THE SIMULATED TIME SERIES.
C      INPUT/OUTPUT VECTOR OF LENGTH 2. ALPHA CONTAINS, WHEN
C      I=1, INPUT MINIMUM SIGNIFICANCE LEVEL OF THE MODEL
C      IN THE EXCLUSIVE RANGE (.1,.1). ON OUTPUT, THE
C      ESTIMATED SIGNIFICANCE LEVEL IS RETURNED.
C      I=2, INPUT VALUE IN THE EXCLUSIVE RANGE (.0,.1)
C      USED FOR COMPUTING  $\text{SIM}(1-\text{ALPHA}(2))$  PERCENT
C      PROBABILITY LIMITS FOR THE FORECASTS.
C      =OUTPUT VECTOR OF LENGTH IPDQ(5)+IPDQ(5) CONTAINING
C      ESTIMATES OF THE AUTOREGRESSIVE PARAMETERS.
C      PMAS =OUTPUT VECTOR OF LENGTH IPDQ(4) CONTAINING ESTIMATES
C      OF THE MOVING AVERAGE PARAMETERS.
C      PMAC =OUTPUT ESTIMATE OF THE TREND FACTOR
C      PWV =OUTPUT ESTIMATE OF THE WHITE NOISE VARIANCE
C      FCST =OUTPUT MATRIX OF DIMENSION 3 BY IPDQ(6),
C      FCST(I,J) = FTH LEAD TIMES J=1,2,...,IPDQ(6),
C      CONTAINS, WHEN
C      I=1, THE WEIGHTS FOR THE WEIGHTED SUM OF SHOCKS
C      THAT GIVES THE FORECAST ERROR
C      I=2, THE FORECASTS
C      I=3, THE CORRESPONDING DEVIATION FROM EACH
C      FORECAST FOR THE 100(1-ALPHA(2)) PERCENT
C      PROBABILITY LIMITS.
C      SIM =OUTPUT MATRIX OF DIMENSION IPDQ(6) BY IPDQ(7)
C      DEFINED ONLY WHEN IPDQ(7) IS POSITIVE. SIM(I,J,J)
C      FOR LEAD TIMES I=1,2,...,IPDQ(6), CONTAINS THE
C      RESULTS OF THE IPDQ(7) SIMULATIONS.
C      TS =FIRST DIVISION OF SIM AS DIMENSIONED IN THE
C      CALLING PROGRAM.
C      WK =WORK AREA LENGTH DEFINED IN THE PROGRAMMING NOTES
C      IER =ERROR PARAMETER.
C      TERMINAL ERROR #1284N
C      N#1, INDICATES A TERMINAL ERROR OCCURRED IN FTMAPS
C      OR FTARPS
C      N#2, INDICATES ONE OF THE INPUT PARAMETERS IPDQ(1)
C      WAS OUT OF RANGE
C      N#3, INDICATES ALPHA(I), I=1,2 WAS NOT IN THE
C      EXCLUSIVE RANGE (0,1).
C      WARNING SIGNAL #22+N
C      N#4, INDICATES THE MODEL DID NOT PASS THE
C      AUTOMATIC SIGNIFICANCE TEST. THE NEED FOR MORE
C      PARAMETERS IS INDICATED.
C      1. THIS ROUTINE IS OBTAINED BY MODIFYING THE IMBL
C      ROUTINE FTIMP.
C      2. WK IS A WORK VECTOR OF LENGTH AT LEAST THE MAXIMUM OF
C      618
C      GTS 9M
C      C      A. 2*IPDQ(2)*42
C      C      B. IPDQ(2)*P*(IPDQ(2)*G+1)
C      C      C. IPDQ(2)*( ((G-3)*Q)/2 )**Q*P*36
C      C      WHERE P,D,Q ARE NUMBER OF PARAMETERS OF THE ARIMA(P,D,Q)
C      C      ----- DIVISION
C      C      7(1)/IPDQ(1),ALPHA(1),DARPS(1),FCST(3,1),
C      C      1 (DIMENSION, MFTAN(60))
C      C      DOUBLE PRECISION SEED
C      C      DATA 10/8.,11/113/15/5/128/28/
C      C      SET OUTPUT PARAMETERS OF IPDQ
C      C      1010 CONTINUE
C      C      N=IPDQ(2)
C      C      IP=IPDQ(3)
C      C      GTS 16M
C      C      GTS 17M
C      C      GTS 18M
C      C      GTS 19M
C      C      GTS 20M
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      WRITE(6,922)
      CALL PATTY2(COV,NV,NNAME,1)
      RETURN
      IF(RU<NE)C(123) GO TO 48
      97 WRITE(6,923)
      WRITE(6,922)
      WRITE(6,923)
      WRITE(6,922)
      98 SMALL=-10.0E+36.0
      SMALL = 10.0E+36.0
      DO 39 I=1,NV
      WRITE(6,912) I
      30 32 I=1,N
      C(I,I)=0.0
      DO 31 K=1,NV
      C(I,K)=C(I,1)*X(I,K)+Z(I,K)
      31 32 C(I,1)=C(I,1)*Q
      DO 33 I=1,N
      WRITE(6,911,N)
      RANKMALL
      DO 38 J=1,N
      T=(C(J,1)-RANK)38+38.36
      36 T=(C(J,2)-99.0)37+38.36
      37 RANKC(J,1)
      N=NJ
      38 CONTINUE
      C(NJ,2)=99.0
      WRITE(6,911) RANK,NJ
      39 CONTINUE
      40 GO TO 18
      40 FORMAT(19H0COMPONENT ANALYSIS)
      902 FORMAT(12H PROBLEM NO.1)
      903 FORMAT(12H CASE ORDERED BY)
      904 FORMAT(12H CORRELATION COEFFICIENT MATRIX)
      905 FORMAT(12H SIZE OF EACH PRINCIPAL COMPONENT SEPARATELY)
      906 FORMAT(12H CASE NO.)
      911 FORMAT(16H COMPONENT NO.13-12H CASE NO.)
      912 FORMAT(16H COMPONENT NO.16H)
      922 FORMAT(25H EIGEN VALUE CHECK MATRIX)
      923 FORMAT(1H0)
      924 FORMAT(1H0 EIGEN VALUE CHECK MATRIX)
      925 FORMAT(1H0 TOTAL VARIANCE)
      936 FORMAT(1H0,*,7F15.2)
      942 FORMAT(12A)
      986 FORMAT(6A18)
      END

      WRITE((6,922))
      CALL PATTY2(COV,NV,NNAME,1)
      RETURN
      END

      FUNCTION ANUMBP(I)
      ENCODE(A1,NUMBP) I
      1 FORMAT(1A)
      RETURN
      END

      SUBROUTINE EIGEN(MVALU,N,M)
      2 EIGEN SUBROUTINE EIGEN FOR BMDBIM
      3 SUBROUTINE EIGEN(MVALU,N,M)
      C

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