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## Higher or Lower Hemoglobin Transfusion Thresholds for Preterm Infants

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## Abstract

**BACKGROUND**—Limited data suggest that higher hemoglobin thresholds for red-cell transfusions may reduce the risk of cognitive delay among extremely-low-birth-weight infants with anemia.

**METHODS**—We performed an open, multicenter trial in which infants with a birth weight of 1000 g or less and a gestational age between 22 weeks 0 days and 28 weeks 6 days were randomly assigned within 48 hours after delivery to receive red-cell transfusions at higher or lower hemoglobin thresholds until 36 weeks of postmenstrual age or discharge, whichever occurred first. The primary outcome was a composite of death or neurodevelopmental impairment (cognitive delay, cerebral palsy, or hearing or vision loss) at 22 to 26 months of age, corrected for prematurity.

**RESULTS**—A total of 1824 infants (mean birth weight, 756 g; mean gestational age, 25.9 weeks) underwent randomization. There was a between-group difference of 1.9 g per deciliter (19 g per liter) in the pretransfusion mean hemoglobin levels throughout the treatment period. Primary outcome data were available for 1692 infants (92.8%). Of 845 infants in the higher-threshold group, 423 (50.1%) died or survived with neurodevelopmental impairment, as compared with 422 of 847 infants (49.8%) in the lower-threshold group (relative risk adjusted for birth-weight stratum and center, 1.00; 95% confidence interval [CI], 0.92 to 1.10;  $P = 0.93$ ). At 2 years, the higher- and lower-threshold groups had similar incidences of death (16.2% and 15.0%, respectively) and neurodevelopmental impairment (39.6% and 40.3%, respectively). At discharge from the hospital, the incidences of survival without severe complications were 28.5% and 30.9%, respectively. Serious adverse events occurred in 22.7% and 21.7%, respectively.

**CONCLUSIONS**—In extremely-low-birth-weight infants, a higher hemoglobin threshold for red-cell transfusion did not improve survival without neurodevelopmental impairment at 22 to 26 months of age, corrected for prematurity. (Funded by the National Heart, Lung, and Blood Institute and others; TOP [ClinicalTrials.gov](https://clinicaltrials.gov) number, [NCT01702805](https://clinicaltrials.gov/ct2/show/study/NCT01702805).)

Packed red-cell transfusions are commonly used to treat low hemoglobin levels in anemic infants in neonatal intensive care units (NICUs).<sup>1</sup> Infants with an extremely-low birth weight (<1000 g) are at high risk for anemia because of immaturity, impaired erythropoiesis, and frequent blood sampling.

Thresholds for transfusion vary<sup>2</sup> because evidence from randomized trials is limited.<sup>3–7</sup> The largest trial to inform our trial protocol (available with the full text of this article at [NEJM.org](https://www.nejm.org)) involved more than 450 neonates. That trial showed no difference between

low and high hemoglobin transfusion thresholds with respect to the primary outcome of clinically significant complications in neonates to 36 weeks of postmenstrual age<sup>3</sup> or the risk of death or severe adverse neurodevelopmental outcomes at 18 to 21 months of age, corrected for prematurity.<sup>6</sup> However, a post hoc analysis suggested that the risk of mild-to-moderate cognitive delay was reduced with higher hemoglobin thresholds.<sup>6</sup>

To further examine the effect of transfusion practice on neurodevelopment in extremely-low-birth-weight infants, we conducted the Transfusion of Prematures (TOP) trial to test the hypothesis that a higher hemoglobin threshold for red-cell transfusions, as compared with a lower threshold, would reduce the incidence of death or neurodevelopmental impairment in infants at 22 to 26 months of age, corrected for prematurity. Both transfusion algorithms used in the trial were consistent with those used in current practice.<sup>2</sup>

## METHODS

### TRIAL DESIGN AND OVERSIGHT

This open, multicenter, randomized, controlled trial was designed by the authors and conducted in 19 centers (41 NICUs) participating in the Neonatal Research Network of the Eunice Kennedy Shriver National Institute of Child Health and Human Development (NICHD), in collaboration with the National Heart, Lung, and Blood Institute (NHLBI). Infants underwent randomization between December 31, 2012, and April 12, 2017, with follow-up through February 3, 2020.

Four of the authors employed by RTI International, the data coordinating center for the Neonatal Research Network, undertook coordination, monitoring, and data collection, storage, management, and analysis. The authors employed by RTI International had full access to all the data in the trial and take responsibility for the integrity of the data and the accuracy of the data analysis. The first five authors and the last author prepared a draft of the manuscript, which was reviewed and approved by the trial subcommittee and all site investigators. The NICHD and NHLBI staff had input into the trial design, conduct, analysis, and drafting of the manuscript. All the authors vouch for the accuracy and completeness of the reporting and for the fidelity of the trial to the trial protocol.

Investigators at each center (listed in Table S1 in the Supplementary Appendix, available at [NEJM.org](https://www.nejm.org)) and RTI International obtained approval of the institutional review board at each participating site and supervised the informed consent process and trial procedures. Written informed consent was obtained from the parent or legal guardian of each child. Independent oversight was provided by the data and safety monitoring committee of the Neonatal Research Network, which included experts in neonatology, bioethics, and biostatistics, as well as a neonatal transfusion specialist appointed for this trial by the NHLBI.

### PARTICIPANTS

Infants with a birth weight of 1000 g or less, a gestational age between 22 weeks 0 days and 28 weeks 6 days, and a postnatal age of 48 hours or less were eligible to participate in the trial. Infants were excluded if they were considered to be nonviable by the attending neonatologist, had cyanotic congenital heart disease, had parents who were opposed to

blood transfusion, had a parent with hemoglobinopathy or congenital anemia, had received a transfusion in utero, had twin-to-twin transfusion syndrome or isoimmune hemolytic disease, or had received a previous red-cell transfusion after the first 6 hours of life. Other exclusion criteria were the likelihood that the infant's family would not be able to return for follow-up assessment at 22 to 26 months, receipt or planned receipt of erythropoietin, and a congenital condition (other than premature birth) adversely affecting life expectancy or neurodevelopment.

## RANDOMIZATION

The infants were randomly assigned in a 1:1 ratio to the higher- or lower-threshold group. Randomization was performed centrally by telephone with the data coordinating center. Randomization was stratified according to birth weight (<750 g or 750 to 1000 g) and trial center and balanced within randomly chosen block sizes of two or four patients. Multiple-birth siblings underwent randomization individually.

## INTERVENTIONS

All routine red-cell transfusions in both groups were guided by algorithms until the infants reached 36 weeks of postmenstrual age or discharge from the NICU of the trial hospital, whichever occurred first (Table S2). Transfusion algorithms were consistent with current practice,<sup>2,8</sup> including practice at participating Neonatal Research Network sites. We adhered to ethical recommendations regarding trials of dose-adjusted therapies.<sup>9</sup> Hemoglobin transfusion thresholds in both groups were determined according to postnatal age (highest in the first week of life, lower in each of the 2 successive weeks, and stable thereafter) and according to the use of respiratory support (a higher threshold when respiratory support was warranted). Respiratory support was defined as mechanical ventilation, continuous positive airway pressure, a fraction of inspired oxygen (F<sub>IO</sub><sub>2</sub>) greater than 0.35, or delivery of oxygen or room air by nasal cannula at a flow of 1 liter per minute or more.

With perfect adherence to these transfusion algorithms throughout the treatment period, the mean between-group difference in the pretransfusion hemoglobin levels was expected to be 2.0 to 2.5 g per deciliter (20 to 25 g per liter). The transfusion volume was 15 ml per kilogram of body weight. All transfusions received by the infants were from ABO-compatible and Rh-compatible donors. The clinicians were permitted to deviate from the applicable algorithm temporarily for reasons specified in the protocol (e.g., if the infant had bleeding or anticipated bleeding during or after surgery, severe hypoxemia, hypotension, or sepsis). All the blood transfusions were recorded and centrally reviewed for consistency with the algorithms; if they were inconsistent, they were deemed to be deviations. All deviations were independently adjudicated as being either justified by the protocol or violations according to agreement of two investigators who were unaware of the treatment assignments.

## Primary Outcome

The primary outcome was a composite of death or neurodevelopmental impairment in infants at 22 to 26 months of age, corrected for prematurity. All the follow-up examiners, including psychologists and psychometrists, were unaware of the treatment assignments.

Neurodevelopmental impairment was defined as one or more of the following components: cognitive delay, moderate or severe cerebral palsy, or severe vision or hearing loss. Cognitive delay was defined as a composite cognitive score of less than 85 (1 SD below the mean of 100) on the Bayley Scales of Infant and Toddler Development, third edition<sup>10</sup>; scores range from 55 to 145, with higher scores indicating better performance. Moderate cerebral palsy was defined as level II or III in the Gross Motor Function Classification System (GMFCS) (levels range from I [mild impairment] to V [most severe impairment]), and severe cerebral palsy as GMFCS levels IV or V.<sup>11</sup> Severe vision loss was defined as a corrected visual acuity in the better eye of less than 20/200. Severe hearing impairment was defined as bilateral hearing loss for which hearing aids or cochlear implants were warranted, according to a hearing assessment conducted by sound field testing or according to testing of auditory brain-stem responses. We prespecified that infants without moderate or severe cerebral palsy and without vision or hearing loss had to complete the cognitive subtest of the Bayley Scales successfully in order to be counted as unimpaired for the composite primary outcome.

### Secondary Outcomes

Table S3 lists all the prespecified secondary outcomes to first discharge home and at 2 years of age. Prespecified secondary outcomes at 22 to 26 months of age included death, neurodevelopmental impairment and its four components, and more detailed analyses of the composite cognitive, language, and motor scores on the Bayley Scales. After a protocol change in 2013, the parents or guardians of the infants completed the Child Behavior Checklist<sup>12</sup> instead of the Brief Infant–Toddler Social and Emotional Assessment in order to comply with revised Neonatal Research Network follow-up procedures.

Prespecified secondary neonatal outcomes included survival to initial hospital discharge without severe complications. These complications included grade 3 or 4 intraventricular hemorrhage, cystic periventricular leukomalacia, or ventriculomegaly diagnosed with the use of clinically performed ultrasonographic examination of the head<sup>13</sup>; stage 3 or greater retinopathy of prematurity or receipt of treatment for that condition<sup>14</sup>; and bronchopulmonary dysplasia diagnosed at 36 weeks of postmenstrual age on the basis of a documented need for supplemental oxygen (defined as an  $FiO_2$  of 0.30 or inability to pass the oxygen reduction test if the  $FiO_2$  was 0.22 to 0.29).<sup>15</sup> Stage 2 or 3 necrotizing enterocolitis was recorded.<sup>16</sup> The infant's head circumference, weight, and length at a postmenstrual age of 36 weeks, postmenstrual age at the last use of caffeine therapy, and the number of packed red-cell transfusions were also secondary outcomes. For each transfusion, it was noted whether the administration was according to the trial protocol.

### STATISTICAL ANALYSIS

Our main objective was to assess the between-group difference in the composite primary outcome of death or neurodevelopmental impairment at 22 to 26 months of age, corrected for prematurity. On the basis of a previous trial,<sup>6</sup> we hypothesized that there would be an absolute between-group difference of 7 percentage points in the incidence of death or neurodevelopmental impairment. Given a 52% incidence of death or neurodevelopmental impairment (unpublished data) in the Neonatal Research Network in 2005–2008 (with entry criteria that were similar to those in this trial), we assumed outcome rates of 53.5% for

the lower-threshold group and 46.5% for the higher-threshold group, centered around a conservative overall event rate of 50% in the two treatment groups.. We estimated that with a sample of 1824 infants and 10% loss to follow-up, the trial would have 80% power to detect an absolute difference of 7 percentage points in the incidence of the primary outcome between the trial groups, at a two-tailed type I error rate of 0.05.

The independent data and safety monitoring committee reviewed the incoming data at four time points during the enrollment period, as prespecified by the trial protocol. Because primary outcome data were available only at 22 to 26 months of follow-up, statistical interim monitoring at four intervals focused on a composite safety outcome of in-hospital death, necrotizing enterocolitis, or adverse findings on ultrasonographic examination of the head. The significance of the interim safety analyses with O'Brien–Fleming boundaries was calculated with a Lan–DeMets spending function to preserve an overall type I error rate of 0.05 for the composite safety outcome.

All the analyses were adjusted for stratification according to birth-weight group and center, and all the infants were evaluated according to treatment assignment. The primary outcome was analyzed with robust Poisson regression to obtain adjusted relative risks and 95% confidence intervals. The consistency of the treatment effect across birth-weight strata and sexes and treatment heterogeneity across the various centers were examined by adding suitable interaction terms to the adjusted models. Binary secondary outcomes, including red-cell transfusion outcomes, were analyzed with the use of robust Poisson regression. Linear regression was used for the number of transfusions received. Other secondary outcomes, such as the time to regain birth weight and time to full enteral feeding, were analyzed with Cox proportional-hazards survival regression, and data were censored for deaths. There was no prespecified plan to adjust for multiplicity of testing for the secondary outcomes, and a P value is reported only for the primary outcome (with a two-sided P value <0.05 considered to indicate statistical significance); all the other analyses are exploratory. Analyses were conducted with the use of SAS software, version 9.4 (SAS Institute).

## RESULTS

### PATIENTS

The composite primary outcome at 22 to 26 months of age could be determined for 1692 of the 1824 infants enrolled (92.8%). (Fig. 1). Selected characteristics of these infants and their mothers are shown in Table 1, along with those of the entire trial cohort, according to treatment group. All the baseline characteristics of the 911 infants who were randomly assigned to the higher-threshold group were similar to those of the 913 infants who were assigned to the lower-threshold group, and these characteristics were similar in the two groups of infants in the final analysis cohort for the primary outcome. Additional characteristics of the infants are provided in Table S4.

### SEPARATION OF THE HEMOGLOBIN LEVELS AND PROTOCOL VIOLATIONS

At the time of randomization, the mean ( $\pm$ SD) hemoglobin levels were similar in the two groups ( $13.8\pm 2.6$  g per deciliter in the higher-threshold group and  $13.7\pm 2.6$  g per deciliter

in the lower-threshold group) (Table 1). Subsequently, the pretransfusion mean hemoglobin levels differed between the groups by 1.9 g per deciliter (19 g per liter) throughout the treatment period ( $P < 0.001$ ) (Fig. 2A). A comparison of all hemoglobin levels obtained throughout the infants' hospital stay confirmed separation between the mean hemoglobin levels in the two groups (Fig. 2B). Of all red-cell transfusions, 3.5% were adjudicated to be protocol violations, and 2.5% of all the transfusions mandated by the respective transfusion algorithms were not administered (Table S5).

## PRIMARY OUTCOME

Complete data for the primary outcome were available for 1692 of 1824 infants (92.8%). Of the 845 infants with data for this outcome who were assigned to the higher transfusion threshold, 423 (50.1%) died or survived with impairment, as compared with 422 of 847 infants (49.8%) assigned to the lower transfusion threshold (adjusted relative risk, 1.00; 95% confidence interval [CI], 0.92 to 1.10;  $P = 0.93$ ). Similar results were obtained in a post hoc sensitivity analysis in which a generalized estimating equation was used to account for any correlation among siblings of a multiple birth (adjusted relative risk, 1.01; 95% CI, 0.92 to 1.10;  $P = 0.88$ ). The incidences of all components of the primary outcome were similar in the two groups (Table 2).

In addition, there was no evidence that the effects of the transfusion strategy on the primary outcome differed according to center, birth-weight group, or sex (Fig. S1). In sensitivity analyses accounting for missing primary outcome data, the results were materially unchanged, regardless of whether all missing outcomes were assumed to be events (adjusted relative risk, 1.01; 95% CI, 0.93 to 1.09) or non-events (adjusted relative risk, 1.00; 95% CI, 0.91 to 1.10).

## SECONDARY OUTCOMES

The incidences of prespecified short-term and long-term secondary outcomes were similar among infants in the higher-threshold group and those in the lower-threshold group (Table 3). The mean ( $\pm$ SD) number of transfusions was  $6.2 \pm 4.3$  and  $4.4 \pm 4.0$ , respectively (mean difference, 1.71; 95% CI, 1.37 to 2.05) (Table 3). A total of 885 of 911 infants (97.1%) in the higher-threshold group and 804 of 913 infants (88.1%) in the lower-threshold group received at least one transfusion (adjusted relative risk, 1.10; 95% CI, 1.08 to 1.13). Additional transfusion data are provided in Table S6. Results for other prespecified and post hoc secondary outcomes are provided in Tables S7 and S8. Serious adverse events occurred in 22.7% of the infants in the higher-threshold group and 21.7% of those in the lower-threshold group (adjusted relative risk, 1.04; 95% CI, 0.88 to 1.23) (Table S9).

## DISCUSSION

Our trial showed that among extremely-low-birth-weight infants, the risk of death or neurodevelopmental impairment at 22 to 26 months of age, corrected for prematurity, was not significantly lower with a higher hemoglobin transfusion threshold level than with a lower hemoglobin transfusion threshold level during the initial hospital course. Although a post hoc analysis of a previous trial had suggested a moderate cognitive benefit at 18 to

21 months of age in infants who were randomly assigned to maintain a higher hemoglobin level,<sup>6</sup> the current, much larger trial showed no evidence to support an improvement in this or other components of the composite primary outcome or in any other clinically important outcome, whether measured during the initial hospital course or at 22 to 26 months of age. These results are consistent with the failure of erythropoietin to improve cognitive or other neurodevelopmental outcomes despite increasing red-cell mass.<sup>18</sup> Our findings are consistent with the results of a similar but smaller European trial that were published while our manuscript was under review.<sup>19</sup>

Although strategies to individualize red-cell transfusions with the use of physiological measures are promising, further testing is needed.<sup>20</sup> Correspondingly, the use of blood transfusion in NICUs remains high. A Canada-wide study conducted in 2010–2012 showed that 82% of infants with a birth weight between 501 and 750 g received red-cell transfusions.<sup>21</sup> As in our trial, several groups in Europe<sup>22</sup> and North America<sup>23</sup> have shown that the use of transfusions decreases when a strict guideline is adopted. In the intervention in the current trial algorithm, we chose hemoglobin levels that would remain within clinically accepted ranges.<sup>2,8,9</sup>

Previous observational data on preterm infants have suggested risks associated with blood transfusions. A large observational cohort study in Brazil involving very-low-birth-weight infants showed an excess cumulative hazard ratio for death in those who received blood transfusions,<sup>24</sup> but this observation may have been explained by the increased severity of illness in these infants.<sup>25</sup> Some observational studies have shown an association between transfusions and necrotizing enterocolitis, but observational data have been inconsistent, and this finding has not been supported by available data from randomized trials.<sup>26–30</sup> One prospective study showed that necrotizing enterocolitis was not associated with transfusion and was more likely to occur when nadir hemoglobin levels decreased to below 8 g per deciliter (80 g per liter) before transfusion.<sup>30</sup> In other observational studies, transfusions have been linked to retinopathy of prematurity,<sup>31</sup> bronchopulmonary dysplasia,<sup>32</sup> and intraventricular hemorrhage.<sup>33</sup> In contrast, other studies have suggested that the risks of hypoxemia and apnea of prematurity are increased among infants who do not receive transfusions.<sup>34,35</sup> We recorded the postmenstrual age at the last use of caffeine therapy as a proxy for apnea, and the results were similar in the two threshold groups. We found no other effects of transfusions on beneficial or adverse outcomes, including stage 2 or 3 necrotizing enterocolitis. Such discrepancies between associations shown in observational studies and our findings underscore the pitfalls of observational studies.<sup>36,37</sup> Although our trial was not powered to address these individual adverse outcomes, it was a large trial comparing high transfusion thresholds with low transfusion thresholds in this vulnerable population.

Our trial has some important limitations. Blinding of the trial intervention was not feasible at the bedside. However, follow-up examiners were unaware of the treatment assignments. For ethical reasons, we could not withhold nonalgorithmic transfusions (i.e., those that were not performed according to the randomly assigned transfusion algorithm), so there was an imbalanced violation rate, with more nonalgorithmic transfusions in the lower-threshold group. This imbalance presumably reflected the unease of some physicians with hemoglobin levels in the lower range. Nonetheless, the incidence of violations was low and did



not preclude good between-group separation in mean hemoglobin levels. Blood banks at different centers did not have uniform practices; it would not have been feasible to manage this pragmatic multicenter trial with the use of a single blood bank. However, this variation is unlikely to have affected outcomes, since randomization was stratified according to center. Moreover, apart from leukoreduction,<sup>38</sup> other blood-banking practices such as transfusing only fresh red cells<sup>39,40</sup> have not been shown to have beneficial effects. The usual clinical variation in blood-banking practices enhances the generalizability of our findings.

In our trial, a higher hemoglobin threshold for transfusion was associated with an increase in the number of transfusions administered. However, it did not improve survival without neurodevelopmental impairment at 22 to 26 months of age among extremely-low-birth-weight infants.

## Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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## APPENDIX

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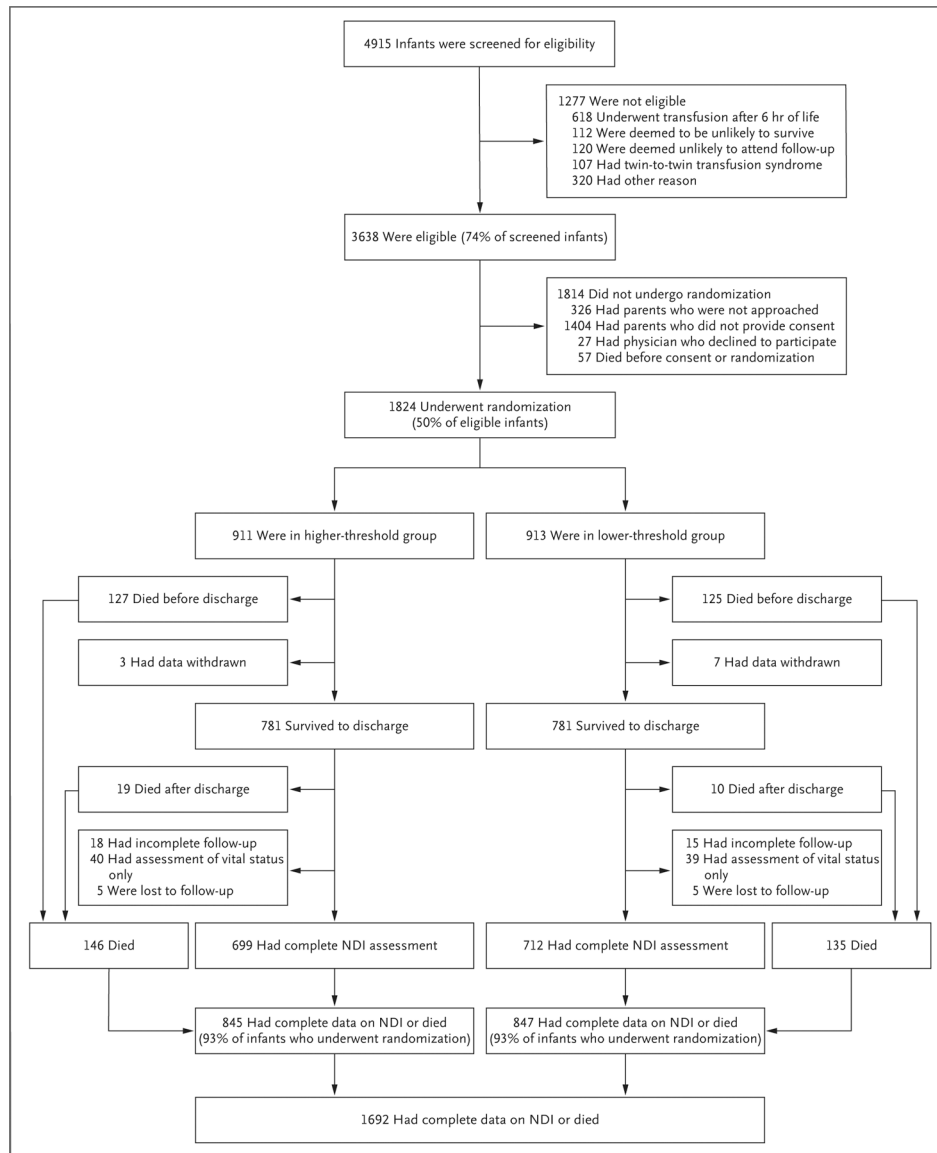
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**Figure 1 (facing page). Screening, Randomization, Intervention, and Follow-up.**

Infants were screened for eligibility only if they met the inclusion criteria of a birth weight of 1000 g or less, a gestational age between 22 weeks 0 days and 28 weeks 6 days, admission to the center neonatal intensive care unit within 48 hours after birth, and the ability to undergo randomization within 48 hours after birth. The assigned intervention was to extend until 36 weeks of postmenstrual age, unless the infant had been transferred or discharged before that age. In the higher-threshold group, 591 infants completed the treatment per protocol up to 36 weeks, and an additional 137 did so with one or more protocol violations, whereas 183 discontinued treatment early because of death, withdrawal, or transfer or discharge. In the lower-threshold group, 562 infants completed the treatment per protocol up to 36 weeks, an additional 186 did so with one or more protocol violations, and 165 discontinued treatment before 36 weeks. In the higher-threshold group, 23 infants withdrew from the intervention early, of whom 3 had parents who did not consent to their

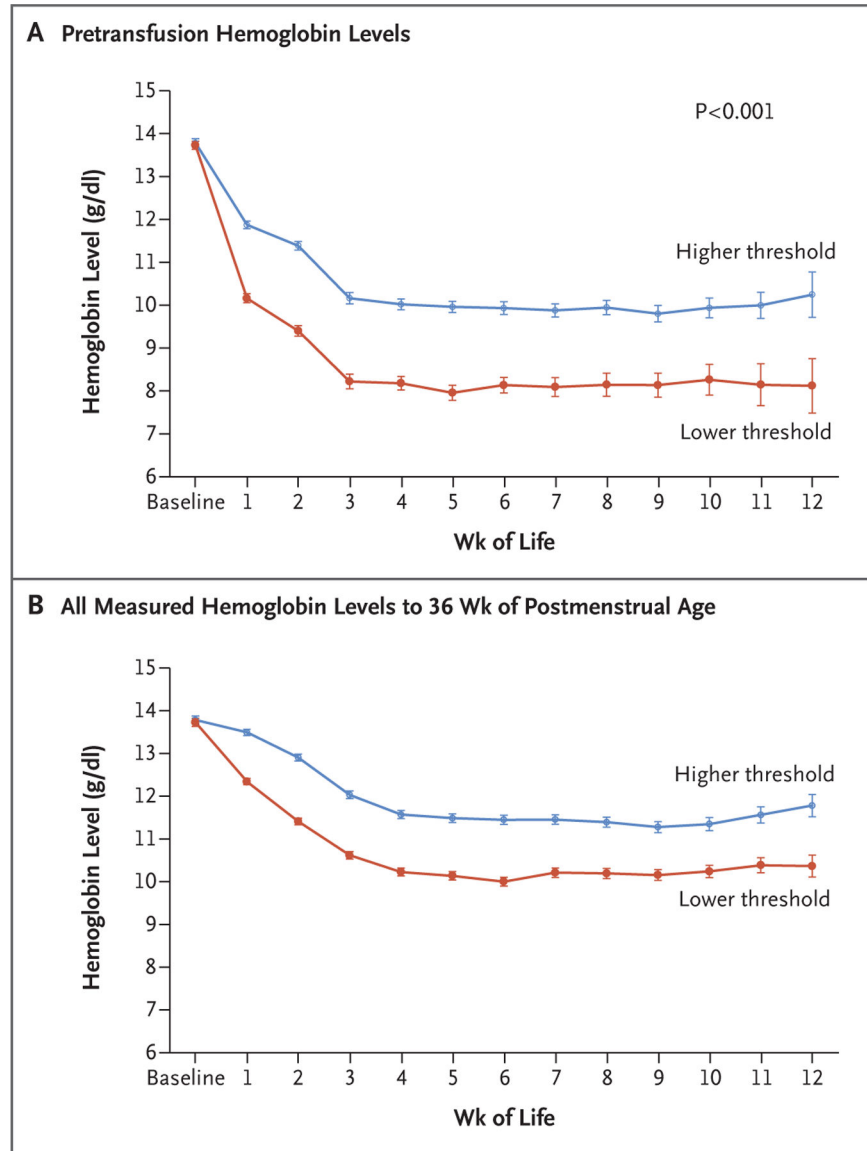
inclusion in subsequent analyses. In the lower-threshold group, 16 infants withdrew from the intervention early, of whom 7 had parents who did not consent to their inclusion in subsequent analyses. Two infants who died before 36 weeks of postmenstrual age also withdrew, but in this figure they were counted as infants who died. Infants with incomplete follow-up included those who presented for the follow-up examination but were missing a key component, usually the cognitive subtest of the Bayley Scales of Infant and Toddler Development assessment (14 infants in the higher-threshold group and 8 infants in the lower-threshold group), and those who did not present for examination but had parents who provided limited questionnaire data (4 and 7 infants, respectively). NDI denotes neurodevelopmental impairment.

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**Figure 2. Separation of Hemoglobin Levels between the Treatment Groups.**

Hemoglobin levels in the higher-threshold and lower-threshold groups were recorded before enrollment and until 36 weeks of postmenstrual age. Values are means and 95% confidence intervals (indicated by I bars), adjusted for infant as a random effect. Hemoglobin tests were performed at clinical discretion and were not dictated by protocol. Panel A shows the hemoglobin levels that prompted a red-cell transfusion. Panel B shows all hemoglobin levels that were measured in the two groups during the treatment period.

**Table 1.**

**Neonatal and Maternal Characteristics at Baseline.\***

Characteristic	Analysis Cohort <sup>†</sup>		All Participants	
	Higher Hemoglobin Threshold (N = 845)	Lower Hemoglobin Threshold (N = 847)	Higher Hemoglobin Threshold (N = 911)	Lower Hemoglobin Threshold (N = 913)
<b>Infant characteristics</b>				
Birth weight — g	754.9±152.7	757.3±148.6	755.2±152.7	757.2±147.7
Gestational age — wk	25.9±1.5	25.9±1.5	25.9±1.5	25.9±1.5
Female sex — no. (%)	441 (52)	428 (51)	488 (54)	462 (51)
Singleton — no./total no. (%)	655/845 (78)	656/847 (77)	708/911 (78)	703/910 (77)
Born at trial hospital — no./total no. (%)	793/845 (94)	784/847 (93)	851/911 (93)	842/910 (93)
Transfusion within 6 hr after birth — no. (%)	47 (6)	36 (4)	47 (5)	40 (4)
Received any antenatal glucocorticoids — no./total no. (%)	766/844 (91)	751/847 (89)	822/910 (90)	808/910 (89)
Cesarean delivery — no./total no. (%)	540/845 (64)	560/843 (66)	585/911 (64)	606/906 (67)
Documented cord milking or delayed clamping — no./total no. (%)	231/841 (27)	208/843 (25)	245/906 (27)	216/909 (24)
Hemoglobin level at randomization — g per deciliter	13.8±2.6	13.7±2.6	13.8±2.6	13.7±2.6
<b>Maternal characteristics — no. (%)</b>				
<b>Race<sup>‡</sup></b>				
Black	387 (46)	366 (43)	423 (46)	395 (43)
White	396 (47)	408 (48)	419 (46)	437 (48)
Other	47 (6)	52 (6)	53 (6)	56 (6)
Data not available	15 (2)	21 (2)	16 (2)	25 (3)
<b>Education</b>				
Less than high school graduate	144 (17)	147 (17)	158 (17)	160 (18)
High school graduate	237 (28)	232 (27)	261 (29)	256 (28)
More than high school graduate	331 (39)	332 (39)	352 (39)	353 (39)
Data not available	133 (16)	136 (16)	140 (15)	144 (16)
Private insurance	271 (32)	272 (32)	283 (31)	284 (31)

\* Plus-minus values are means ±SD(s).

<sup>†</sup>The analysis cohort comprised infants with complete data for the primary outcome.



Race was reported by the biologic mother on the worksheet for the child's birth certificate.

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Primary Composite Outcome and Components of the Primary Composite Outcome at 2 Years.\*

Table 2.

Outcome	Higher Hemoglobin Threshold (N = 845)		Lower Hemoglobin Threshold (N = 847)	Adjusted Relative Risk (95% CI)	P Value
	no. of infants	total no. (%)			
Primary outcome: death or neurodevelopmental impairment	423	845 (50.1)	422	1.00 (0.92–1.10)	0.93
Components of primary outcome					
Death †	146	903 (16.2)	135	1.07 (0.87–1.32)	
Neurodevelopmental impairment	277	699 (39.6)	287	1.00 (0.88–1.13)	
Cognitive delay ‡	269	695 (38.7)	270	1.04 (0.91–1.18)	
Moderate or severe cerebral palsy §	48	711 (6.8)	55	0.87 (0.60–1.26)	
Severe vision impairment ¶	5	713 (0.7)	6	0.83 (0.25–2.76) ¶¶	
Severe hearing impairment	14	710 (2.0)	25	0.56 (0.29–1.07) ¶¶	

\* The relative risk was adjusted for birth-weight and center stratum. The P value is reported only for the composite primary outcome. For the secondary outcomes, 95% confidence intervals (CIs) were not adjusted for multiplicity and should not be used to infer definitive treatment effects.

† Infants with incomplete follow-up but with known vital status were included in these numbers. Five infants in each group had unknown vital status at the start of the 2-year follow-up window.

‡ Cognitive delay was defined as a composite cognitive score of less than 85 (1 SD below the mean of 100) on the Bayley Scales of Infant and Toddler Development (BSID-III), third edition; scores range from 55 to 145, with higher scores indicating better performance.<sup>10</sup>

§ Moderate or severe cerebral palsy was defined as a Gross Motor Function Classification System level II or higher (levels range from I [mild impairment] to V [most severe impairment]). Level II denotes moderate cerebral palsy with a limited ability to walk. Levels III to V indicate increasing severity of gross motor impairment.

¶ Trial centers with low incidences of blindness or hearing impairment were pooled with their nearest geographic center before estimation of adjusted relative risks.

Table 3.

Prespecified Secondary Outcomes to Hospital Discharge and at 2 Years.\*

Outcome	Higher Hemoglobin Threshold	Lower Hemoglobin Threshold	Adjusted Relative Risk, Hazard Ratio, or Mean Difference (95% CI) <sup>†</sup>
<b>Outcomes to hospital discharge</b>			
Survival to discharge without severe complications — no./total no. (%) <sup>‡</sup>	257/901 (28.5)	274/888 (30.9)	0.93 (0.81 to 1.06)
Bronchopulmonary dysplasia at 36 wk postmenstrual age — no./total no. (%) <sup>§</sup>	469/795 (59.0)	453/805 (56.3)	1.04 (0.96 to 1.13)
Retinopathy of prematurity stage 3 or treatment for that condition received — no./total no. (%) <sup>¶</sup>	157/797 (19.7)	137/797 (17.2)	1.14 (0.94 to 1.39)
Grade 3 or 4 intraventricular hemorrhage, cystic periventricular leukomalacia, or ventriculomegaly diagnosed on ultrasonographic examination — no./total no. (%)	146/855 (17.1)	154/859 (17.9)	0.94 (0.77 to 1.16)
Necrotizing enterocolitis, Bell's stage 2 — no./total no. (%) <sup>  </sup>	91/907 (10.0)	95/906 (10.5)	0.95 (0.73 to 1.25)
No. of transfusions per infant	6.2±4.3	4.4±4.0	1.71 (1.37 to 2.05)
<b>Anthropometric measures**</b>			
Weight for age			
No. of infants	769	774	
Change in z score	-1.2±0.8	-1.3±0.8	0.04 (-0.04 to 0.11)
Length for age			
No. of infants	715	715	
Change in z score	-1.9±1.0	-1.9±0.9	0.07 (-0.01 to 0.16)
Head circumference for age			
No. of infants	754	766	
Change in z score	-1.4±1.0	-1.4±1.0	-0.01 (-0.10 to 0.08)
<b>Postmenstrual age at final tracheal extubation in infants who were intubated</b>			
No. of infants	796	804	
Wk	30.1±3.4	30.2±3.3	-0.11 (-0.43 to 0.21)
<b>Postmenstrual age at final caffeine dose in infants who received caffeine treatment</b>			
No. of infants	882	887	
Wk	33.8±3.0	34.0±2.8	-0.19 (-0.45 to 0.07)
<b>Length of hospital stay<sup>††</sup></b>			
No. of infants	908	906	

Outcome	Higher Hemoglobin Threshold	Lower Hemoglobin Threshold	Adjusted Relative Risk, Hazard Ratio, or Mean Difference (95% CI) <sup>†</sup>
Median days (IQR)	96 (72 to 129)	97 (75 to 127)	-1.25 (-6.96 to 4.48)
Time to full enteral feeding <sup>††</sup>			
No. of infants	808	824	
Median days (IQR)	19.5 (14 to 29)	19.0 (15 to 30)	0.96 (0.87 to 1.05)
<b>Outcomes at 2 yr — no./total no. (%)</b>			
Severe cerebral palsy <sup>§§¶¶</sup>	20/710 (2.8)	11/720 (1.5)	1.85 (0.90 to 3.82)
Hydrocephalus shunt <sup>¶¶</sup>	20/717 (2.8)	22/728 (3.0)	0.92 (0.51 to 1.67)
Microcephaly <sup>¶¶¶</sup>	61/700 (8.7)	52/710 (7.3)	1.17 (0.83 to 1.66)
Seizure disorder <sup>****</sup>	42/714 (5.9)	41/726 (5.6)	1.04 (0.68 to 1.57)
Respiratory disease necessitating readmission before follow-up	248/715 (34.7)	230/726 (31.7)	1.09 (0.95 to 1.26)
Developmental delay assessed with BSID-III <sup>††††</sup>			
Composite language score <85	355/671 (52.9)	368/691 (53.3)	1.01 (0.91 to 1.11)
Composite motor score <85	255/678 (37.6)	280/695 (40.3)	0.96 (0.84 to 1.09)
Composite cognitive score <70	88/695 (12.7)	96/712 (13.5)	0.96 (0.74 to 1.25)
Composite language score <70	164/671 (24.4)	163/691 (23.6)	1.06 (0.88 to 1.27)
Composite motor score <70	87/678 (12.8)	99/695 (14.2)	0.91 (0.70 to 1.18)

\* Plus-minus values are means ±SD. Outcomes to hospital discharge are reported for 908 infants in the higher-threshold group and 906 infants in the lower-threshold group. These numbers were calculated as the number of infants in each group who underwent randomization (911 and 913, respectively), minus the number of infants in each group (3 and 7, respectively) for whom consent for data collection was withdrawn. Outcomes at 2 years are reported for 699 infants in the higher-threshold group and 712 infants in the lower-threshold group (some of the denominators shown are larger than the number of patients who had a complete follow-up assessment because partial data from incomplete assessments were included). Estimates are adjusted for birth-weight and center stratum. IQR denotes interquartile range.

<sup>†</sup> For categorical outcomes, adjusted differences are adjusted relative risks, with the lower-threshold group as the reference. For time to full enteral feeding, differences are adjusted hazard ratios; for the remaining continuous outcomes, they are adjusted mean differences. The 95% confidence intervals are not adjusted for multiplicity and should not be used to infer definitive treatment effects.

<sup>††</sup> Severe complications were defined as bronchopulmonary dysplasia, retinopathy of prematurity of stage 3 or higher or for which treatment was warranted, and an adverse finding on ultrasonographic examination of the head.

<sup>§</sup> Bronchopulmonary dysplasia was diagnosed on the basis of the need for supplemental oxygen after a standardized oxygen reduction test at 36 weeks of postmenstrual age.

<sup>¶</sup> Retinopathy of prematurity was recorded for infants who underwent a retinal examination before discharge from the neonatal intensive care unit.

<sup>¶¶</sup> Bell's stages range from 1 to 3, with higher stages indicating greater severity of disease.

<sup>\*\*\*\*</sup> Z scores are derived from Olsen growth curves.<sup>17</sup>

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<sup>††</sup>Length of stay was up to hospital discharge or death, whichever occurred first.

<sup>††</sup>A total of 808 of 907 infants (89.1%) in the higher-threshold group and 824 of 906 infants (90.9%) in the lower-threshold group attained full enteral feeding during the period of observation. The hazard ratio may be interpreted as the odds of attaining full enteral feeding faster at any point in time. Deaths were considered to be a censoring event.

<sup>§§</sup>Severe cerebral palsy was defined as Gross Motor Function Classification System (GMFCS) levels IV or V on an ordinal scale on which levels range from I (mild impairment) to V (most severe impairment).

<sup>¶¶</sup>The trial centers with low incidences of this condition were pooled with their nearest geographic center.

<sup>|||</sup>Microcephaly was defined as a head circumference-for-age z score of less than -2, according to growth curves developed by the World Health Organization (WHO) (WHO Multicentre Growth Reference Study Group, Geneva; [www.who.int/toolkits/child-growth-standards/head-circumference-for-age](http://www.who.int/toolkits/child-growth-standards/head-circumference-for-age)) in infants in whom age was corrected for prematurity.

<sup>\*\*\*</sup>A seizure disorder was defined as a report of having one or more seizures since discharge or of regular use of anticonvulsants or seizure medications.

<sup>†††</sup>BSID-III composite language and motor scores range from 40 to 160, and BSID-III composite cognitive scores range from 55 to 145, with higher scores indicating better performance.<sup>10</sup> Composite BSID-III scores of less than 85 are less than 1 SD below the mean of 100. Composite BSID-III scores of less than 70 are less than 2 SD below the mean of 100.