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Spatial variability of soil respiration (Rs) and its controls are subjected to strong seasonality in an even-aged European beech (Fagus sylvatica L.) stand

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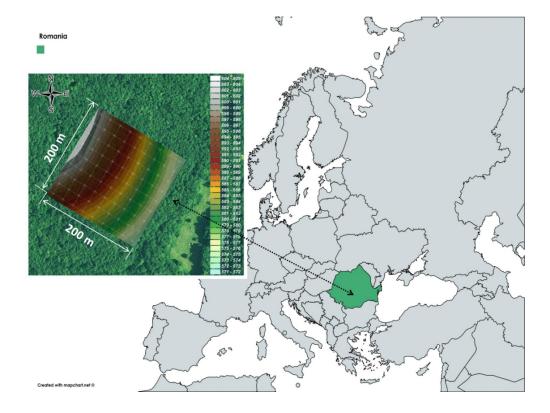
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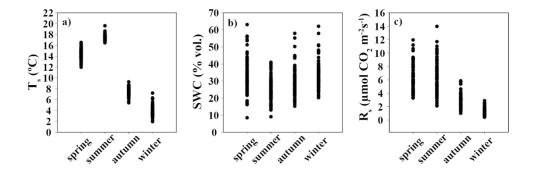
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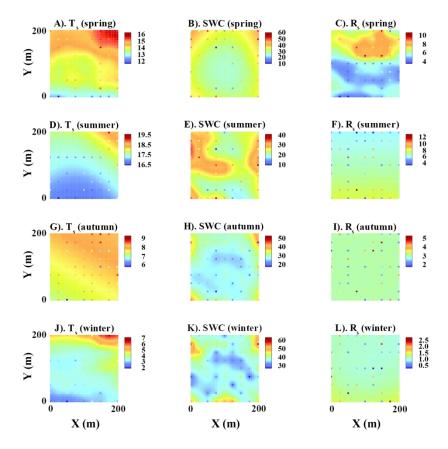
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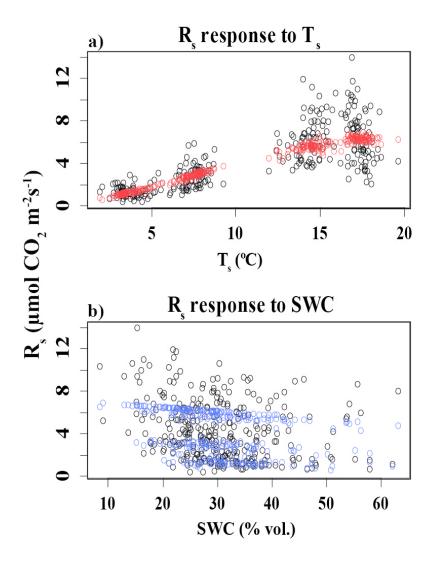




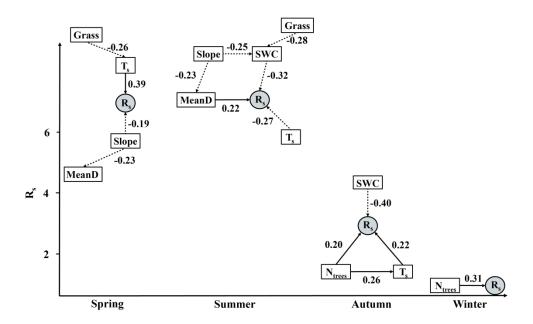
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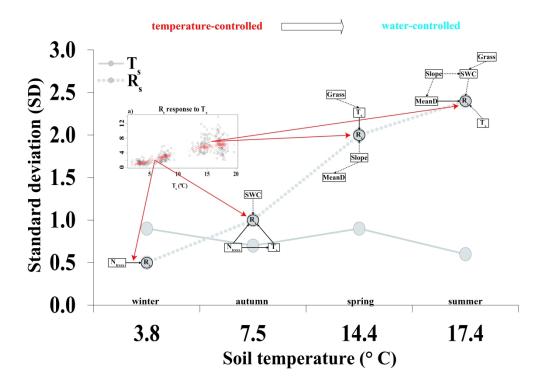
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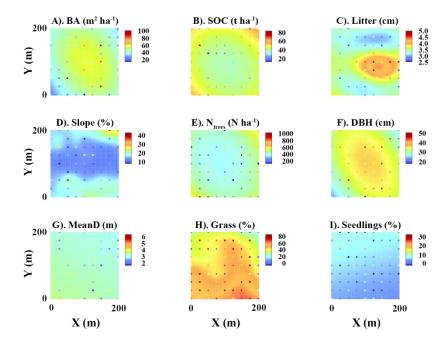
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| 1 | Spatial variability of soil respiration (R_s) and its controls are subjected to strong |
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| 2 | seasonality in an even-aged European beech (Fagus sylvatica L.) stand |
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| 4 | Running Title: Space-Time R _s variability in European beech stand |
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| 31 | Keywords: | European | beech, | even-aged | stand, | micro-topography, | seasonality, | soil |
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microclimate, soil respiration, spatial variability

Abbreviations: R_s, soil respiration; **T**_s, soil temperature; **SWC**, soil water content; **BA**, basal area of the European beech trees surrounding the 81 measurement points; **SOC**, soil organic carbon content; **Litter**, thickness of the litter layer; **Slope**, micro-topography of the terrain within the study stand; **N**_{trees}, the count of all the surrounding European beech trees around each of the 81 measurement points; **DBH**, average diameter at breast height (i.e., > 6 cm) of the European beech trees surrounding the 81 measurement points; **MeanD**, mean distance from the European beech trees to the 81 measurement points; **Grass**, percentage of the soil surface covered by grass; **Seedlings**, percentage of the soil surface covered by tree seedlings.

Abstract

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Uncertainties arising from the so far poorly explained spatial variability of soil respiration (R_s) remain large. This is partly due to the limited understanding on how actually spatially variable R_s is, but also on how environmental controls determine R_s's spatial variability and how these controls vary in time (e.g., seasonally). Our study was designed to deepen into the complexity of R_s's spatial variability in a European beech even-aged stand, covering both phenologicaly and climatically contrasting periods (spring, summer, autumn, winter). Although we studied a relatively homogenous stand, we found a large spatial variability of R_s (coefficients of variation > 30%) characterized by strong seasonality. This large spatial variability of R_s suggests that even in relatively homogenous stands there is a large potential source of error when estimating R_s. This was also reflected by the sampling effort needed to obtain seasonal robust estimates of R_s, which may actually require a number of samples above that used in R_s studies. We further postulate that the effect of seasonality on the spatial variability and environmental controls of R_s was determined by the seasonal shifts of its microclimatic controls: during winter, low temperatures constrain plant and soil metabolic activities and hence reduce R_s variability (temperature-controlled processes), while during summer, water demand by vegetation and changes in water availability due to the micro-topography of the terrain (i.e., slope) increase R_s variability (water-controlled processes). This study provides novel information on the spatiotemporal variability of R_s and deepens into the seasonality of its environmental controls and the architecture of their causal-effect relationships controlling R_s's spatial variability. Our study further shows that improving current estimates of R_s at local and regional levels might be necessary in order to reduce uncertainties and improve CO₂ estimates at larger spatial scales.

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- ❖ The spatial variability of soil respiration (R_s) and its environmental controls vary seasonally
- ❖ Seasonal shifts from temperature- to water-controlled processes determine R_s's spatial 70 variability 71
- * Besides microclimate, slope and grass cover explain the spatio-temporal variability of 72 R_{s} 73
- ❖ An intense sampling effort is needed to obtaining robust R_s estimates even in 74 homogenous forests 75

1. Introduction

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Soil respiration (R_s), i.e., the production and subsequent emission of carbon dioxide (CO₂) from the soil to the atmosphere, is one of the key processes contributing to the global terrestrial carbon (C) balance/budget. R_s is mainly produced by biological sources from the aerobic respiration of decomposers (i.e., heterotrophic respiration), as well as by plant roots and associated microorganisms living in the rhizosphere (i.e., autotrophic respiration) (Rodeghiero & Cescatti, 2008), but also by non-biological chemical oxidation reactions of C in organic matter, although at lower rates in this latter case (Raich & Schlesinger, 1992). Globally, the R_s emissions amount to a total of almost 80 PgC y⁻¹, being the second largest C flux after CO₂ uptake by plants (Raich & Tufekciogul, 2000), which means more than half of an ecosystem's total CO₂ emissions come from R_s (Barba et al., 2018; Curiel Yuste et al., 2005; Janssens et al., 2001). However, R_s is also probably the least well understood part of the C budget at global terrestrial ecosystems' level, based primarily on the fact that the large spatio-temporal variability that characterizes this large flux requires of a substantial monitoring effort at different scales and hence, of a large investment in instrumentation for its correct monitoring (Bond-Lamberty & Thomson, 2010). Therefore, and despite the large critical mass of studies performed to understand it (Bond-Lamberty & Thomson, 2010), our knowledge on the mechanisms controlling this large flux remains very limited (e.g., Barba et al., 2013; Curiel Yuste et al., 2019). Hence, there is still a need for studies designed to explore the spatiotemporal variability of R_s in order to be able to calibrate models and improve predictions of soil biological CO₂ emissions in a changing environment.

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The magnitude of R_s-related CO₂ emissions varies in time and space depending on multiple drivers. A critical mass of studies has been designed to understand how the temporal variability

of R_s relates to different environmental factors such as soil temperature (T_s; e.g., Chen et al., 2014; Davidson et al., 1998; Epron et al., 2004a), soil water content (SWC; e.g., Davidson et al., 2000; Oishi et al., 2013; Poblador et al., 2017), wind (e.g., Sánchez-Cañete et al., 2016), or the photosynthetic activity of the plants (e.g., Bahn et al., 2009; Curiel Yuste et al., 2005; Davidson et al., 1998). Nevertheless, it is important to highlight the discrepancy between the large number of studies undertaken to understand the large, but predominantly explained variability in time (generally seasonal) of soil CO₂ fluxes (see for instance Bond-Lamberty & Thomson, 2010) and the relatively few studies undertaken to understand the enormous, but largely unexplained spatial variability of this very same flux. Several studies have proposed different factors that define local conditions as controls of the spatial variability of R_s at the mesoscale (scale of m). Most of these studies agree on the important role of variables such as: i). the spatial variability of soil moisture (Barba et al., 2013; Kosugi et al., 2007; Poblador et al., 2017); ii). the structure of the overstorey plant community (Barba et al., 2013; Epron et al., 2004b; Law et al., 2001; Saiz et al., 2006; Søe & Buchmann, 2005); iii). variables directly related to the structure of the aboveground plant community, such as leaf production (Oishi et al., 2013), root density or biomass (Knohl et al., 2008), microbial biomass, and litter thickness (Hanson et al., 1993); iv). the quantity and quality of soil organic matter (Rayment & Jarvis, 2000); or v). the C/N ratio and bulk density of the top soil (Khomik et al., 2006; Ngao et al., 2012; Saiz et al., 2006). Other topographical aspects, such as the slope and the position within the landscape, have been however less studied although their contribution to explain the spatial variability of R_s might also be critical (Arias-Navarro et al., 2017; Berryman et al., 2015; Brito et al., 2010; Hanson et al., 1993; Riveros-Iregui et al., 2012;). All studies, nevertheless, conclude that our capacity to predict the spatial variability of R_s and its environmental controls remains largely insufficient (e.g., Allaire et al., 2012).

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The environmental controls of the spatial variability of R_s may also vary temporally, though, to the best of our knowledge, only few studies have been designed to deepen in this potential temporal axis of the spatial variation of R_s (Epron et al., 2004b; Khomik et al., 2006; Kosugi et al., 2007; Saiz et al., 2006; Shi et al., 2016; Søe & Buchmann, 2005). The complexity of the spatial variability of R_s can vary seasonally (Riveros-Iregui et al., 2012; Shi et al., 2016; Søe & Buchmann, 2005) specially because different environmental drivers may differently influence R_s depending on the season. For instance, the influence of water availability on the spatial patterns of R_s at the landscape-scale can exhibit a bidirectional behaviour, R_s being more sensitive to water availability during dry periods or in highly drained areas than during wetter periods or in low drainage areas (Riveros-Iregui et al., 2012). Likewise, the biomass and respiration of the autotrophic (roots and rhizosphere microorganism) and heterotrophic (microbial activity) components of R_s may vary in space and time depending on the phenological state of the vegetation and its nutrient and water demands (Barba et al., 2013; Søe & Buchmann, 2005). For this reason, understanding the drivers controlling the spatial variability of R_s at different temporal scales may help us to improve and modulate the sampling effort needed in order to obtain confident estimates of R_s. This also means that obtaining reliable integrative measures of R_s would require different sampling efforts throughout the year. It is, therefore, important to understand this seasonally-dependent complexity if we want to improve our knowledge on the sampling effort needed to get accurate and costly efficient estimates of R_s (Barba et al., 2013; Herbst et al., 2009; Rayment & Jarvis, 2000; Rodeghiero & Cescatti, 2008).

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We studied the spatio-temporal variability of soil respiration (R_s) in a 4.0 ha (i.e., 200 m x 200 m) European beech (*Fagus sylvatica* L.) even-aged stand. Specifically, we focused on understanding the potential seasonal (i.e., spring, summer, autumn, winter) variations of R_s and

its environmental controls to: (1) determine the magnitude of their spatial variability and the sampling effort needed per each season (i.e., spring, summer, autumn, winter) to obtain robust average estimates of R_s; and (2) identify the main environmental controls and the architecture of their potential causal-effect relationships controlling the spatial variability of R_s along the seasons (i.e., during phenologically and climatically contrasted periods of the year). We hypothesized that, given the generally large influence of the aboveground plant distribution in explaining the spatial variability of R_s, the spatial variability of R_s will be low in our European beech even-aged study stand where trees are homogeneously distributed (H1). However, we also hypothesized that, along with other already well-studied and known factors, other factors (i.e., more spatially variable at stand level), such as the micro-topography of the terrain (i.e., slope) or the spatial distribution of the grass cover, will also play an important, indirect control over R_s due to their influence on the spatial variability of the soil water content (SWC) (H2). Finally, we also hypothesized that the predictive power of the different environmental controls of the spatial variability of R_s will vary throughout the year depending on the environmental constrains that act on R_s within a given season (e.g., soil temperature, T_s; or soil water content, SWC) (H3).

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2. Materials and Methods

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2.1. Study site and stand

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The study site (i.e., forest) is located in the central-southern part of Romania, in Mihaesti (Arges county; 45°05'11.8019"N, 25°03'58.0428"E), at an altitude of 570 m a.s.l. (Figure 1). This forest is largely dominated by European beech, although other tree species may also be found: hornbeam (*Carpinus betulus* L.), sessile oak (*Quercus petraea* (Matt.) Liebl.), or sweet cherry

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(Prunus avium L.). The density within the whole forest located in Mihaesti is of 504 trees ha⁻¹, with a total volume of 502 m³ ha⁻¹, and a basal area of 33 m² ha⁻¹ (Mihaesti Forest Management Plan). Within this forest, we focused on a 4.0 ha European beech study stand (200 m x 200 m) (Figure 1). The European beech trees within the study stand are mainly adult and dominant (i.e., canopy level). According to the Mihaesti Forest Management Plan, most individuals within the study stand have an estimated age of ~ 85 years, which allows us to consider this study stand as being even-aged. The area where our study stand is located is characterized by a temperate continental climate, with a mean annual precipitation of ~ 875.21 mm, and a mean air temperature of ~ 6.31 °C, respectively (estimates calculated for the 1901 – 2019 period; CRU TS v.4; Harris et al., 2020). The mean annual precipitation for 2016 and 2017 (i.e., the years when our measurements were performed; see below) was of ~ 991.80 mm and ~ 959.60 mm, respectively. As for the mean annual air temperature, it was of ~ 7.59 °C in 2016 and of 7.54 °C in 2017 (CRU TS v.4; Harris et al., 2020). The soils are Eutric Cambisols (clay loam) covered with mull type humus, developed on a sandstone with marls parental material (Florea & Munteanu, 2012). The slope within the study stand is smooth and there are no important differences regarding the altitude between the upper part of the study stand and the lower part of the study stand (Figure 1, small panel). Mean pH values range from 4.8 (0-10 cm soil depth) to 5.2 (11-20 cm soil depth) (WTW pH330i; WTW GmbH, Weilheim, Germany).

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2.2. Field soil respiration (R_s) and microclimatic factors measurements

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The 4.0 ha selected study stand was divided into regular 25 m x 25 m squares (Figure 1, small panel). Soil respiration (R_s) measurements were then performed at each of the four corners of each of the 25 m x 25 m squares, resulting thus on a total of 81 measurement points. R_s measurements were all performed using a Portable Infrared Gas Analyzer (IRGA) connected to

a soil respiration standard chamber (EGM-4 and SRC-1; PP Systems, Amesbury, MA, USA). The soil respiration chamber covered a soil surface area of 78 cm² and an enclosed volume of 1171 cm³. Since some studies have shown a clear correlation between insertion depth, the amount of cut roots, and the lost soil effluxes (Silvola et al., 1996; Wang et al., 2005), no collars were inserted in the soil (Arias-Navaro et al., 2017; Epron et al., 2004b; Hanson et al., 1993; Maestre & Cortina, 2003; Poblador et al., 2017). Instead, we followed a similar procedure to the one described by Epron et al., 2004b and we inserted the edge of the respiration chamber to a depth of 1 cm into the soil, including the litter layer. Nevertheless, this was done only after firstly removing the herbaceous layer in order to avoid potential confounding effects of the vegetation on R_s measurements. Furthermore, to avoid potential gas leaks due to the shallow insertion of the respiration chamber (1 cm into the soil) with respect to a relatively thick lowdensity litter layer (average 3.3 cm; Table 1), the respiration chamber was strongly pressed against the soil (i.e., with the help of one hand) over the whole time measurements were performed. Final R_s values were estimated for 120 seconds based on the linear increase of the CO₂ concentration within the soil respiration chamber (i.e., a closed dynamic system). Soil CO₂ efflux measurements were always performed between 9 a.m. and 5 p.m. Additionally, the CO₂ effluxes were never measured during rainy days. Specifically, in case of heavy rains (i.e., > 15 mm), field R_s measurements were postponed 36 h to avoid the "Birch effect" (Birch, 1958).

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Simultaneously to the field R_s measurements, microclimatic measurements (i.e., soil temperature and the volumetric soil water content) were also performed at the same 81 measurement points. Specifically, soil temperature (T_s) was measured at 5 cm soil depth using the STP-2 Soil Temperature Probe that was attached to the IRGA (PP Systems, Amesbury, MA, USA). As for the volumetric soil water content (SWC), this variable was measured at 20 cm soil depth using the TDR 300 soil moisture meter (Spectrum Technologies, Inc., Plainfield, IL,

USA). All field measurements (i.e., R_s, T_s, and SWC) spanned over a period of one complete year and thus over the four seasons: spring (May 2016), summer (August 2016), autumn (November 2016), and winter (February 2017). Within each of the 4 seasons and at each of the 81 measurement points, we performed 3 independent measurements for each of the 3 variables (i.e., R_s, T_s, and SWC) and then averaged their corresponding values. In order to systematically perform R_s, T_s, and SWC measurements at exactly the same locations within the study stand, we marked the 81 measurement points with wood sticks that were maintained in their positions over the whole study period. Due to the large number of measurement points (i.e., 81) and thus to the considerable field effort and logistics that were needed, R_s, T_s, and SWC measurements were always performed during 2 consecutive days during each season.

2.3. Forest structural and soil variables and the micro-topography of the terrain

At each of the 81 measurement points, soil samples were also collected to determine the soil organic carbon (SOC) content. All soil samples were collected in February 2017 after all seasonal measurements (i.e., R_s, T_s, and SWC) were finished. Soil sampling was performed using a metallic cylinder (5 cm diameter, and 20 cm depth) and consisted in extracting one soil core at each of the 81 measurement points. SOC of the upper 20 cm of the soil profile was determined through the dry combustion method using a CHNS organic elemental microanalyser (TruSpec Micro CHNS elemental analyser, LECO, New York, USA).

The thickness of the litter layer (hereinafter referred to as "litter" to simplify) was used as a proxy of litter biomass, which could not be measured due to logistics. The litter, at each of the 81 measurement points, was measured only once during the 2016 summer, two weeks before the R_s , T_s , and SWC measurements started. Although, we acknowledge the fact that it would

have been better to measure the litter layer over the year (i.e., seasons), this was not possible due to logistics. Instead, we assumed that the place where there was more accumulated litter (i.e., at some point) would be the same place where more litter usually falls and the opposite for the places where there was less accumulated litter. Accordingly, the litter depth would be basically stable over the year (i.e., seasons). Simultaneously to the litter measurements, the micro-topography of the terrain (hereinafter referred to as "slope" to simplify), at each of the 81 measurement points, was also measured.

In order to account for the impact of the surrounding vegetation on our field measurements (i.e., within a radius of 7 m around each of the 81 measurement points), we counted all the surrounding European beech trees (N_{trees}) and we measured their diameter at breast height (DBH; at standard 1.3 m above from the ground) and their distance to the 81 sampling points. The 7 m radius was established considering the average crown diameter of the European beech trees found within the 4.0 ha study stand (Mihaesti Forest Management Plan). The DBH of the trees was measured using a calliper (Haglöf, Sweden), only European beech trees with a DBH > 6 cm being finally considered for this study. The measured distances were used to calculate the mean distances (MeanD) from surrounding European beech trees to the 81 measurement points. In order to estimate the basal area (BA; m² ha⁻¹) of all European beech trees with a DBH > 6 cm, we calculated the sum of all their cross-sectional areas at breast height. Finally, within the same radius of 7 m around each of the 81 measurement points, we also estimated the percentage (%) of the soil surface covered by grass and the percentage (%) of the soil surface covered by all tree seedlings. These estimations were done visually and agreed between several observers for data consistency.

2.4. Statistical analyses

We used different statistics (i.e., mean, M; standard deviation, SD; and relative variability, RV) to estimate relative rates of spatial variability of forest structural and soil variables (i.e., BA, SOC, litter, N_{trees}, DBH, MeanD, % of grass, and % of seedlings) and of the micro-topography of the terrain (i.e., slope). We used the same statistics (i.e., mean, M; standard deviation, SD; and relative variability, RV) plus the absolute amplitude (A; defined as the difference between maximum and minimum values) to estimate relative and absolute rates of spatial variability of the microclimate (i.e., T_s and SWC) and soil respiration (R_s) variables. As most studies give the coefficient of variation (CV), we also calculated this statistic (i.e., expressed as a percentage) for the R_s variable alone and used it to compare our results with those published in previous studies. For the T_s, SWC, and R_s variables, all the above mentioned statistics were calculated considering both the four seasons separately and the four seasons combined (i.e., annual). The relative variability statistic (i.e., RV) was calculated following Lewontin 1966 and Webster 2001 as the standard deviation of the logarithms (i.e., log10 in our case) of measurements. Relative variability allows thus to compare variations between different groups of observations (Lewontin 1966; Webster 2001).

In order to calculate the minimum number of measurements (i.e., N) needed to obtain robust estimates of R_s for each season (i.e., spring, summer, autumn, and winter), we used the following power equation (Davidson et al., 2002):

$$N = [(t \times s)/(range/2)]^2 \qquad equation 1$$

where, t is represented by the critical value of the t-distribution (two-tailed test) for a given confidence level (99, 95, and 90%, respectively) and for 80 degrees of freedom; s is the standard

deviation of all spatially independent R_s measurements per each season (i.e., spring, summer, autumn, and winter); and range is the width of the desired interval around the mean of the R_s measurements of each season (i.e., spring, summer, autumn, and winter) in which a smaller sample mean is expected to fall (i.e., error limit of 10%, 20%, and 30% of the R_s measurements mean per each season).

We run geostatistical analyses (i.e., experimental (semi-) variograms and theoretical

variograms) to determine the spatial autocorrelation of: i). the T_s, SWC, and R_s variables,

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separately for each season (i.e., spring, summer, autumn, and winter); ii). all forest structural 310 311 and soil variables (i.e., basal area, BA; soil organic carbon content, SOC; thickness of the litter layer, litter; number of trees surrounding the 81 measurement points, N_{trees}; diameter at breast 312 height, DBH; mean distance from surrounding European beech trees to the 81 measurement 313 314 points, MeanD; percentages of grass and seedlings cover around the 81 measurement points); and iii). the micro-topography of the terrain (i.e., slope). Specifically, the experimental (semi-) variograms (i.e., binned) were performed using the "variog" function available from the "geoR" 316 R package (Ribeiro et al., 2020) based on classical estimators. Further on, the theoretical 317 variograms were performed using the "likfit" function available from the "geoR" R package 318 319 (Ribeiro et al., 2020). These analyses were run considering the restricted maximum likelihood (REML) parameter estimation, different trends (i.e., the mean part of the model; constant, first order polynomial, and second order polynomial), and functions (i.e., models for the correlation 321 function; matern, exponential, Gaussian, spherical, circular, cubic, wave, powered exponential, 322

Cauchy, gneiting, and pure nugget). A total of 693 models were run. The selection of the best

models for each of the analysed variables was based on AIC (Akaike Information Criteria). The

selected models were then used to perform ordinary kriging using the "krige.conv" function

from the "geoR" R package (Ribeiro et al., 2020) was finally used to visualize the spatial prediction based on the fixed covariance parameters. All selected models were validated using the "xvalid" function available from the "geoR" R package (Ribeiro et al., 2020).

To evaluate the microclimatic controls of the spatio-temporal variability of R_s , we run multiple regression functions using the "nls" function available from the "MASS" R package (Venables & Ripley, 2020). These functions were designed according to literature (i.e., Vicca et al., 2014). Briefly, different models were designed to represent the independent controls of soil temperature (T_s) and soil water content (SWC) on R_s , but also taking into account potential unimodal responses of R_s to both microclimatic factors. R_s data was log transformed prior to analyses as it did not meet the normality assumption.

$$\begin{split} \log(R_s) \sim a + b*T_s & \textit{equation 2} \\ \log(R_s) \sim a + b*SWC & \textit{equation 3} \\ \log(R_s) \sim a + b*T_s + c*T_s^2 & \textit{equation 4} \\ \log(R_s) \sim a + b*SWC + c*SWC^2 & \textit{equation 5} \\ \log(R_s) \sim a + b*T_s + c*SWC & \textit{equation 6} \\ \log(R_s) \sim a + b*T_s + c*SWC + d*T_s^2 & \textit{equation 7} \\ \log(R_s) \sim a + b*T_s + c*SWC + d*SWC^2 & \textit{equation 8} \\ \end{split}$$

Where, a, b, c, and d letters stand for coefficients of the multiple regression functions. The covariance and multicollinearity between T_s and SWC were examined prior to analyses using the Variance Inflation Factor (VIF). Since the VIF was lower than 2, both microclimatic variables could be used within the same model (Zuur et al., 2010). The selection of the best model was based on the AIC (Akaike Information Criterion). For these analyses, the T_s , SWC, and R_s

variables were combined over the four seasons (i.e., spring, summer, autumn, and winter). All these analyses were conducted based on the assumption that the residuals of the models were normally distributed (p > 0.05) and independent. As the assumption of independence may be justified when data collection has been performed based on a probability sampling design (e.g., de Gruijter et al., 2006) and our data collection has been performed based on a grid sampling design, we acknowledge the fact that the spatial auto-correlation between data at nearby measurement points might impact some of the obtained results.

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To describe potential complex causal-effect relationships that might determine the spatiotemporal variability patterns of R_s, we run Structural Equation Models (SEMs). SEMs analyses allowed to test for the direct and indirect effects of all our measured in the field variables (i.e., microclimatic variables, forest structural and soil variables, and the micro-topography of the terrain) on seasonal (i.e., spring, summer, autumn, and winter) R_s. SEMs analyses were carried out using the "psem" function available from the "piecewiseSEM" R package (Lefcheck, 2016). To harmonize the results of the SEMs with those obtained from the multiple regression functions analyses, R_s was logarithmically transformed prior to analyses. Separated SEMs were built for each of the four seasons (i.e., spring, summer, autumn, and winter). All SEMs were designed based on hypotheses supported on simple univariate correlations between the different microclimatic (i.e., soil temperature, T_s; soil water content, SWC), micro-topography of the terrain (i.e., slope), soil (i.e., soil organic carbon content, SOC; and litter), and forest structural (i.e., basal area, BA; number of trees, N_{trees}; diameter at breast height, DBH; mean distance, MeanD; % of grass; and % of seedlings) variables (i.e., predictor variables). Furthermore, their potential complex causal-effect relationships, that might determine the spatio-temporal variability of R_s, were also considered. To test the goodness of fit of the SEMs, the Fisher's C statistic was calculated. The Fisher's C statistic follows a chi-squared distribution and tests if

the model fits the data (p > 0.05) or not (p < 0.05). Several SEMs were run, the selection of the best one being based on the AIC (Akaike Information Criteria) (Lefcheck, 2016).

All statistical analyses were carried out in R (v. 4.0.0, R Core Team, 2020). Statistical relationships were considered significant at p < 0.05.

3. Results

3.1. The spatial variability of the forest structural and soil variables around the 81 soil respiration (R_s) measurement points

We found a relatively low spatial variability (i.e., expressed here as relative variability, RV; Table 1) for forest structural variables such as DBH (RV = 0.1) and MeanD (RV = 0.1) (Table 1, Figure S1). Similar results were obtained also for soil variables such as SOC (RV = 0.1) and litter (RV = 0.1) (Table 1, Figure S1). On the other hand, the spatial variability of the understory vegetation (i.e., % of grass and % of seedlings) showed the highest values and was higher relative to the spatial variability of the overstorey vegetation (i.e., BA and N_{trees}) (Table 1, Figure S1). High spatial variability values were found also for the slope (RV = 0.3) (Table 1, Figure S1).

| G: .: | BA | SOC | Litter | Slope | N _{trees} | DBH | MeanD | Grass | Seedlings |
|------------|-----------|-----------------------|--------|-------|-----------------------|------|-------|-------|-----------|
| Statistics | (m² ha-1) | (t ha ⁻¹) | (cm) | (%) | (N ha ⁻¹) | (cm) | (m) | (%) | (%) |
| M | 47.3 | 44.3 | 3.3 | 12.2 | 443 | 32.8 | 3.5 | 57.5 | 9.1 |
| SD | 16.9 | 12.9 | 0.7 | 8.3 | 200 | 6.3 | 0.8 | 31.1 | 8.6 |
| RV | 0.2 | 0.1 | 0.1 | 0.3 | 0.2 | 0.1 | 0.1 | 0.4 | 0.3 |

Table 1. Mean (M), standard deviation (SD), and relative variability (RV; following Lewontin 1966 and Webster 2001) values of forest structural, soil, and the micro-topography of the terrain variables. *Where*, **BA**, basal area of the European beech trees surrounding the 81 measurement points; **SOC**, soil organic carbon content; **Litter**, thickness of the litter layer; **Slope**, microtopography of the terrain within the study stand; **N**_{trees}, the count of all the surrounding European beech trees around each of the 81 measurement points; **DBH**, average diameter at breast height (i.e., > 6 cm) of the European beech trees surrounding the 81 measurement points; **MeanD**, mean distance from the European beech trees to the 81 measurement points; **Grass**, percentage of the soil surface covered by grass; **Seedlings**, percentage of the soil surface covered by tree seedlings.

3.2. Soil respiration (R_s) sampling effort needed per season

The calculation of the minimum number of measurements needed to obtain robust estimates of R_s for each season (i.e., spring, summer, autumn, and winter) showed how sensitive *equation 1* was to both the error limit (i.e., 10%, 20%, and 30%) and the confidence interval (90%, 95%, and 99%) (Table S1). Accordinly, the calculated sampling effort varied within and among seasons depending on the error limit and the confidence interval. Specifically, the sampling effort varied more strongly within seasons than among them (Table S1). As for the sampling effort needed from one season to another to obtain robust estimates of R_s , differences were not so strong especially between summer, autumn, and winter (Table S1). Spring was the season when less minimum R_s measurements seem to be needed to obtain robust estimates of R_s (Table S1) no matter the confidence level and the error limit (Table S1). On the other hand, winter was found to be the season when more sampling effort seems to be needed to obtain robust estimates of R_s no matter the confidence level and the error limit (Table S1). The sampling effort for the

summer season seems to be quite close to the winter one, while the sampling effort for the autumn season was found to be slightly lower (i.e., in between the spring and summer), no matter the confidence level and the error limit (Table S1).

3.3. Seasonal and spatial variability of soil respiration (R_s) , soil temperature (T_s) , and soil water content (SWC)

As expected, both microclimatic variables (i.e., T_s and SWC) experienced very different seasonal patterns during the study period (Table 2, Figure 2). On one hand, T_s experienced large seasonal changes, peaking during summer and reaching its minimums during winter (Table 2, Figure 2a). On the other hand, values of SWC experienced less seasonality, reaching its minimums during summer, but being very stable and similar for the rest of the year (Table 2, Figure 2b). Seasonality of R_s followed a pattern similar to that observed for T_s , peaking during both spring and summer, and reaching its minimums during winter (Table 2, Figure 2c).

The spatial variability of the microclimatic variables was also markedly different (Table 2, Figure 3). Specifically, the spatial variability of T_s was generally smaller (RV values ranging from 0.01 to 0.10), within the range of 3.1 to 5.3 °C of amplitude (Table 2, Figure 3A, D, G, J), than the spatial variability of SWC (RV values ranging from 0.10 to 0.14), within the range of 31.9 to 54.6 % vol. of amplitude (Table 2, Figure 3B, E, H, K). R_s was the variable with the highest spatial variability (RV values ranging from 0.13 to 0.17), within the range of 2.5 to 11.9 μ mol CO_2 m⁻² s⁻¹ of amplitude (Table 2, Figure 3C, F, I, L) and coefficients of variation above 30% (Table 2). As concerning the standard deviation of the mean values, both T_s and SWC showed seasonal changes, summer being the season that showed the lowest values (Table 2). Standard deviation values for R_s also showed seasonal changes, being higher in spring and

summer and smaller in winter (Table 2), a pattern that followed the increase of the R_s rates (Figure 2).

| | | T_s | | | | SWC | | | | $\mathbf{R}_{\mathbf{s}}$ | | | | | |
|--------|------|-------|------|------|----------|------|------|------|---|---------------------------|------|------|------|--|--|
| Period | (°C) | | | | (% vol.) | | | | (μmol CO ₂ m ⁻² s ⁻¹) | | | | | | |
| | M | SD | RV | A | M | SD | RV | A | M | SD | RV | CV | A | | |
| Spring | 14.4 | 0.9 | 0.03 | 4.6 | 33.0 | 10.2 | 0.14 | 54.6 | 6.3 | 2.0 | 0.13 | 31.1 | 8.7 | | |
| Summer | 17.4 | 0.6 | 0.01 | 3.1 | 24.9 | 6.6 | 0.12 | 31.9 | 6.4 | 2.4 | 0.17 | 37.6 | 11.9 | | |
| Autumn | 7.5 | 0.7 | 0.04 | 3.9 | 29.6 | 8.1 | 0.11 | 42.7 | 2.8 | 1.0 | 0.15 | 34.9 | 4.8 | | |
| Winter | 3.8 | 0.9 | 0.10 | 5.3 | 33.9 | 8.0 | 0.10 | 41.9 | 1.3 | 0.5 | 0.16 | 36.4 | 2.5 | | |
| Annual | 10.8 | 5.5 | 0.27 | 17.7 | 30.4 | 9.0 | 0.13 | 54.6 | 4.2 | 2.7 | 0.33 | 65.1 | 13.6 | | |

Table 2. Mean (M), standard deviation (SD), relative variability (RV; following Lewontin 1966 and Webster 2001), Coefficient of Variation (CV; only for the R_s variable) expressed as a percentage, and absolute amplitude (A; defined as the difference between maximum and minimum values) values of soil microclimate (i.e., T_s, soil temperature; and SWC, soil water content) and soil respiration (R_s) measurements. The above mentioned statistics have been calculated both at the seasonal level (i.e., spring, summer, autumn, and winter) and over the four seasons combined (i.e., annual).

3.4. The spatio-temporal variability of soil respiration (Rs)

The model that best explained the microclimatic controls over the spatio-temporal variability of R_s was the model that considered unimodal responses of R_s to T_s and a negative linear response of R_s to SWC (i.e., equation 7) (Table S2, Figure 4). Specifically, the unimodal effect of T_s on R_s translated into a seasonal sensitivity of R_s to T_s : i.e., R_s response to T_s was stronger at lower T_s values (i.e., corresponding to autumn and winter seasons) than at higher T_s values

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(i.e., corresponding to spring and summer seasons) (Figure 4a) when R_s reached its peak (Table 2) and its response to the seasonal changes of T_s was low. The overall effect of SWC over the spatio-temporal variability of R_s was negative, this response being evidenced by the negative slope of this relationship (Figure 4b). This negative effect was further on highlighted by the overall relationship that was found between SWC and R_s at the spatial scale (i.e., results of the SEM analyses), where higher values of SWC were generally associated with low R_s rates (see Figures 4b and 5). Hence, the SWC effect on R_s was mainly driven by the spatially net negative effect of SWC on R_s .

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Both multiple regression functions and SEMs agreed that the microclimatic variables (T_s and SWC; Figures 4 and 5) were, overall, the most important factors controlling the variability of R_s in our 4.0 ha study stand. Nevertheless, SEMs further showed how the architecture of the potential causal-effect relationships controlling R_s's spatial variability increased in complexity (Figure 5) during spring and summer, when the spatial variability of R_s was maximal (see SD and A in Table 2 and Figure 2C). Specifically, SEMs highlighted how during the summer season, when the spatial variability of R_s was the highest (Table 2), the number of variables ultimately involved in controlling the spatio-temporal variability of R_s were also high compared to, e.g. winter or autumn (Figure 5). Also, the predictive capacity of the spatio-temporal variability of R_s varied seasonally (Table 3), the coefficient of determination (R²) ranging from 0.10 in winter to 0.29 in autumn. SEMs showed how both the forest structural (N_{trees}, MeanD, and % of grass cover) and the micro-topography of the terrain (i.e., slope) variables may strongly influence, directly and indirectly the spatio-temporal variability of R_s (Figure 5). Specifically, in spring and summer, the % of grass cover was negatively associated with T_s and SWC, which on their turn exerted a further positive and negative, respectively, influence on the spatio-temporal variability of R_s (Figure 5). N_{trees} instead, always showed a direct and positive

effect on the spatio-temporal variability of R_s during the coldest seasons (i.e., autumn and winter; Figure 5). The slope (i.e., the micro-topography of the terrain within the study stand) was negatively related with MeanD in spring and summer, this relationship being especially important during summer when MeanD exerted some control (i.e., positive relationship) over R_s (Figure 5).



| Season | Response | Predictor | Estimate | SE | df | p-value | SRW | Response | R² | n | Fisher's C | df | p-value |
|----------|---------------------------|--------------------|----------|--------|----|---------|-------|----------------|------|----|------------|----|---------|
| Spring | R_s | T_s | 0.0583 | 0.0154 | 78 | 0.0003 | 0.39 | R _s | 0.17 | 81 | 10.193 | 10 | 0.424 |
| | R_s | Slope | -0.0031 | 0.0017 | 78 | 0.0672 | -0.19 | MeanD | 0.05 | | | | |
| Spring | MeanD | Slope | -0.0213 | 0.0103 | 79 | 0.0413 | -0.23 | T _s | 0.07 | | | | |
| | T_s | Grass | -0.0076 | 0.0031 | 79 | 0.0170 | -0.26 | | | | | | |
| | $\mathbf{R}_{\mathbf{s}}$ | SWC | -0.0086 | 0.0028 | 76 | 0.0030 | -0.32 | R _s | 0.24 | 80 | 12.3 | 12 | 0.422 |
| | R_s | T_s | -0.0836 | 0.0320 | 76 | 0.0109 | -0.27 | SWC | 0.11 | | | | |
| C | R_s | MeanD | 0.0471 | 0.0227 | 76 | 0.0412 | 0.22 | MeanD | 0.05 | | | | |
| Summer | SWC | Slope | -0.1890 | 0.0829 | 77 | 0.0253 | -0.25 | | | | | | |
| | SWC | Grass | -0.0571 | 0.0226 | 77 | 0.0134 | -0.28 | | | | | | |
| | MeanD | Slope | -0.0213 | 0.0103 | 79 | 0.0413 | -0.23 | | | | | | |
| | $\mathbf{R}_{\mathbf{s}}$ | T_s | 0.0504 | 0.0233 | 77 | 0.0331 | 0.22 | R _s | 0.29 | 81 | 1.581 | 2 | 0.454 |
| A | R_s | SWC | -0.0076 | 0.0018 | 77 | 0.0001 | -0.40 | T _s | 0.07 | | | | |
| Autumn | R_s | N _{trees} | 0.0002 | 0.0001 | 77 | 0.0486 | 0.20 | | | | | | |
| | T _s | N _{trees} | 0.0008 | 0.0004 | 79 | 0.0190 | 0.26 | | | | | | |
| Winter | R _s | N _{trees} | 0.0003 | 0.0001 | 79 | 0.0045 | 0.31 | R _s | 0.10 | 81 | 0 | 0 | 1 |

Table 3. Statistics of the Structural Equation Models (SEMs) analyses showing causal-effect relationships that determine the spatio-temporal variability of soil respiration (R_s). Only significant (p < 0.05) and marginally significant (p < 0.1) causal relationships are given. Where, T_s , soil temperature; SWC, soil water content; Slope, micro-topography of the terrain within the study stand; N_{trees} , the count of all surrounding European beech trees around each of the 81 measurement points; MeanD, mean distance from the European beech trees to the 81 measurement points; Grass, percentage of the soil surface covered by grass; SE, standard error; df, degrees of freedom; SRW, Standardized Regression Weights; R², the coefficient of determination; \mathbf{n} , sampling size; **Fisher's C statistic**, follows a chi-squared distribution and tests if the model fits the data (p > 0.05) or not (p < 0.05). The right hand part of the table shows the statistics of the best models representing the spatio-temporal variability of R_s during each of the four different seasons (i.e., spring, summer, autumn, and winter). r).

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4. Discussion

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We here report high soil respiration (R_s) coefficient of variation values (i.e., CV, ranging from 31.1 in spring to 37.6 in summer; see Table 2) in an even-aged European beech study stand located in the central-southern part of Romania (Mihaesti, Arges county). These values are higher or comparable to other CV values mentioned in previous studies (e.g., Barba et al., 2013; Epron et al., 2004b; Kosugi et al., 2007; Ngao et al., 2012; Shi et al., 2016; Stoyan et al., 2000), although caution should be taken when comparing CV values among studies as they might also vary depending on the considered spatial scales (e.g., Darenova & Čater, 2020; Ngao et al., 2012). Nevertheless, independent of this consideration, the high CV values we obtained here refute our first hypotheses (H1). In fact, the magnitude of the spatial variability of R_s during the warmest seasons (i.e., spring and summer) was comparable to the overall annual variability of R_s (see SD and A values in Table 2), which reinforces the idea of the large, though generally neglected, impact of R_s's spatial variability on estimates of soil CO₂ effluxes, even in homogenous ecosystems such as the European beech even-aged study stand that we considered here. The calculated large sampling effort needed to obtain robust estimates of R_s for any given season (being even larger in winter, see Table S1) further highlights the importance of the spatial variability of R_s as a potential source of uncertainty on local and global CO₂ estimates and that should be taken into account. This is of upmost importance especially now, when the number of studies dedicated to scale up CO₂ observations from local to global levels is growing. Accordingly, our study suggests that obtaining robust estimates of R_s at the local level may require of more intense spatial sampling efforts, than those generally carried out for logistical reasons, in order to address and diagnose uncertainties on CO₂ estimates at the global level (e.g., Jian et al., 2018; Warner et al., 2019).

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The large spatio-temporal variability of R_s was strongly and directly determined by soil microclimatic conditions (T_s and SWC; Figures 4 and 5). Nevertheless, as hypothesized (H2), less considered variables such as the forest structural ones (i.e., grass, MeanD, N_{trees}) or the micro-topography of the terrain (i.e., slope), proved to have a determinant, direct or indirect, effect on the observed spatio-temporal variability of R_s. In the case of the slope and the grass cover variables, they both showed further tight relationships with soil microclimatic conditions (i.e., T_s and SWC). These relationships were found to be significant in spring and summer (Figure 5), the two seasons when R_s values peaked (Table 2). Specifically, grass cover modulated the T_s variable in spring, with an indirect effect over R_s. In summer instead, when SWC usually registers low values and the competition for water and nutrients between the heterotrophic communities and the vegetation is high (Villegas et al., 2010), grass cover modulated the SWC availability, with an indirect effect over R_s. Our results highlight thus the importance of seldom considered variables, such as the micro-topography of the terrain (e.g., Arias-Navarro et al., 2017) and the vegetation (e.g., Søe & Buchmann, 2005), in R_s studies, as they may actually substantially impact, either directly or indirectly, the spatio-temporal variability of R_s. Instead, in our European beech even-aged study stand, we found no significant effects of variables generally well associated with the spatial variability of R_s, such as the litter (e.g., Epron et al., 2004b; Katayama et al., 2009; Saiz et al., 2006) or the soil organic carbon content (e.g., Søe & Buchmann, 2005). Although the fact that litter thickness was only measured once (i.e., during the 2016 summer; cf. 2.3. section), and thus may have generated a certain source of noise in our models (since the litter generally accumulates in autumn in deciduous-dominated forests), we assumed that our summer measurements contain very valuable information on the long-term spatial patterns of litter accumulation on the soil, and therefore, valuable information on where litter can have a greater impact on the spatial variability of soil processes in the long term. We are further aware of the limitations of our

study, since other variables, such as soil compaction (Schwen et al., 2015) or the spatial distribution of the root biomass (Søe & Buchmann, 2005), that have not been measured, may have also helped to explain the observed spatial patterns of R_s. However, we expect that effects of the spatial variability of soil compaction on, e.g. water infiltration or CO₂ diffusivity (e.g., Schwen et al., 2015), will not be as high as in more intensively used stands since our study stand has not undergone any forestry intervention during the last 85 years and has no livestock load (according to the Mihaesti Forest Management Plan). On the other hand, and given the logistical inability to obtain estimates of the spatial distribution of root biomass, our exhaustive characterization of the distribution of trees and understorey (e.g., grass and seedlings) around the 81 measurement points emerged as a good proxy highly associated with the distribution of roots, assuming that proximity to vegetation is closely associated with root density in the soil (Søe & Buchmann, 2005).

Our results further emphasized the importance of understanding the temporal (i.e., seasonal) changes in the magnitude and controls of spatial variability of R_s . This variability could be especially important in temperate areas where microclimatic conditions (i.e., T_s and SWC), soil CO_2 effluxes, and vegetation activity may vary dramatically throughout the year (e.g., Curiel Yuste et al., 2005). Indeed, and also as hypothesized (H3), our results indicated that, along with the observed seasonal variability in the magnitude of R_s , the variables that control the spatial variability of R_s were also subjected to strong seasonality. The architecture of the causal-effect relationships controlling R_s 's spatial variability varied between the four seasons and showed an increased complexity during spring and summer, while in autumn and especially in winter these relationships were much simpler (Figure 5). In line with these findings, spring and summer were also the seasons when the highest R_s values were registered, as expected for temperate ecosystems (e.g., Knohl et al., 2008; Saiz et al., 2006; Shi et al., 2016; Søe & Buchmann, 2005).

These high R_s values coincided thus with the warmest temperatures of the year, with the peak in plant and soil biological activity, and with the highest variability of R_s in absolute terms. These results are of upmost importance as they highlight the fact that in order to obtain robust estimates of R_s -CO₂ derived emissions and to have a deeper understanding on the R_s variability, both spatial and temporal R_s controlling processes need to be taken into account.

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Based on our results, we postulate that in this even-aged European beech study stand, the observed spatio-temporal changes and controls of the R_s respond to a seasonal shift that goes from temperature-controlled (i.e., winter and autumn) to water-controlled (i.e., spring and summer) processes. Figure 6 shows a conceptual framework, based on our results, that illustrates this shift in time. During cold periods, when the seasonal variability of R_s was, as expected in a temperature forest (e.g., Curiel Yuste et al. 2003), strongly limited by T_s (Figure 4a), the spatial variability in R_s followed the low spatial variability of T_s (represented as standard deviation of the mean R_s or T_s in Figure 6), resulting in low spatial variability of R_s (Table 2, Figure 6). Most factors had insignificant effects over the spatial variability of R_s during the winter season, when only N_{trees} (i.e., the number of European beech trees surrounding the 81 measurement points) showed a positive relationship with R_s (Figure 5). These results might be related with a larger autotrophic respiration contribution to R_s during winter when European beech trees are able to maintain part of their fine root biomass alive (e.g., Büttner and Leuschner, 1994; Zwetsloot et al., 2019). A similar result was found for the autumn season, when N_{trees} also showed a positive relationship with R_s (Figure 5). During warm periods (i.e., summer), when the soil metabolic activity is at its peak (reflected in higher rates of R_s; Table 2, Figure 2), the increase in temperature and vegetation activity increases the demand for SWC (evapotranspiration), which then becomes a limiting factor for R_s. Although our SWC measurements were too deep (i.e., 20 cm soil depth) to capture this increasing water control

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(i.e., our model could not capture a positive effect of SWC on R_s; see Figure 4b), this seasonal water limitation of R_s was evidenced by the low sensitivity to temperature that R_s experienced during the warmer periods (Figure 4a), which is the shape typically observed when R_s responds to a shift that goes from temperature-controlled to water-controlled processes (e.g., Curiel-Yuste et al., 2003, 2005; Davidson & Janssens, 2006). This shift towards R_s's spatial variability being water-controlled resulted in an increase in R_s variability, which subsequently become more spatially variable than temperature (reflected in increased standard deviation values of R_s with respect to T_s; see Figure 6). The shift towards a water-limited R_s system that generated spatial variability of R_s (Figure 6) also increased the complexity of R_s controls (Figures 5 and 6). This is because the increase in vegetation activity triggered a higher competition for water, as evidenced by, e.g. the strong negative influence of tree proximity (i.e., positive effect of MeanD in summer on R_s; Figure 5) or the strong negative effect of the grass cover over SWC during summer (Figure 5). Hence, the evaporative demand of the vegetation (i.e., MeanD and grass cover) exerted direct and indirect controls over the spatial variability of R_s during dry, warm, and phenologically active periods (i.e., spring and summer), contributing to an increase in the spatial variability of R_s. The slope (i.e., the micro-topography of the terrain) was another variable that contributed, directly and indirectly, to the increase in the spatial variability of R_s during warmer periods characterized by higher water demand (Figure 5). Slope may have large impacts over water availability and water balances by creating spatial variability in e.g. the incidence of solar radiation at the floor level and water run-off (Berryman et al., 2015; Riveros-Iregui et al., 2012), with further consequences on soil CO₂ effluxes, even across short distances (Arias-Navarro et al., 2017). It is likely that, at our study stand, the spatial distribution of the slopes captured the spatial variability of SWC during drier periods (i.e., summer) better than our own SWC measurements taken at 20 cm depth. This is because during periods of high water demand (i.e., summer), SWC decreases very fast in the uppermost layer of the soil, where most

of both autotrophic and heterotrophic activities concentrate (Curiel Yuste et al., 2003, 2005), whereas at 20 cm depth SWC remains above the volumetric content thresholds at which SWC limits Rs, as stated by the fact that no positive relationship was found between SWC and $R_{\rm s}$ (Figures 4b and 5).

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5. Conclusions

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We here highlight the fact that the spatial variability of R_s proves to be high even in a relatively homogenous even-aged European beech study stand of 4.0 ha. Accordingly, our estimates regarding the sampling effort needed to obtain robust estimates of R_s further suggest that most studies to date might have probably underestimated the sampling effort needed to obtain accurate spatial estimates of R_s throughout the year. Our study further shows that the spatial variability of R_s, varied significantly throughout the year, peaking in spring and summer and being low in winter, coinciding thus with the seasonal variability in the absolute magnitude of R_s. We here postulate that in this European beech-dominated even-aged study stand, the observed large seasonal changes in the magnitude and controls of the spatial variability of R_s respond to a seasonal shift that goes from temperature-controlled (i.e., winter and autumn) to water-controlled (i.e., spring and summer) processes. This is because when temperatures and water demands are high, the evaporative demand of both the overstorey but also the understorey vegetation, as well as the micro-topography of the terrain (i.e., slope), generate spatial complexity in soil R_s. During winter, temperature limits processes and prevents most other factors from spatially influencing R_s. In conclusion, obtaining robust, accurate estimates of R_sderived CO₂ effluxes, may profit from: (1) a deeper understanding of how the spatial patterns of R_s varies across seasons, e.g., understanding when processes shift from being controlled by temperature (i.e., winter and autumn) to being controlled by water (i.e., spring and summer);

and (2) a deeper understanding on how, when, and where, factors such as the micro-topography of the terrain or the plant-plant and the plant-soil competition for water may contribute to this spatial variability of R_s. In line with our findings, it would be interesting to test in future at which extent this observed trends apply to other types of ecosystems or if they may also be extrapolated to latitudinal and/or altitudinal gradients, i.e., whether R_s's spatial complexity may increase considering gradients that go from temperature-limited (e.g., temperate) to water-limited (e.g., arid and semi-arid) systems, or from topographically simple (e.g., valleys) to topographically more complex (e.g., mountains) systems.

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Figure captions:

Figure 1. Map indicating the location of Romania within Europe and the location of the study site in the central-southern part of Romania (i.e., Mihaesti, Arges county). The small panel shows the sampling design: a 4.0 ha (i.e., 200 m x 200 m) even-aged European beech stand divided into 25 m x 25 m squares. The scale that appears on the right size of the small panel indicates the altitude (m a.s.l.) gradient within the study stand. Figure 2. Seasonal (i.e., spring, summer, autumn, and winter) patterns of: a) soil temperature (T_s) ; b) soil water content (SWC); and c) soil respiration (R_s) . Figure 3. Spatial prediction based on the fixed covariance parameters generated by performing geostatistical analyses on the seasonal (i.e., spring, summer, autumn, and winter) spatial distribution of soil temperature (T_s), soil water content (SWC), and soil respiration (R_s). **Figure 4**. Representation of the best model (Table S2) that explained the microclimatic controls (soil temperature, T_s; and soil water content, SWC) over the spatio-temporal variability of soil respiration (R_s): a) R_s response to T_s; and b) R_s response to SWC. Black opened dots represent the row data, while red (i.e., R_s response to T_s model) and blue (i.e., R_s response to SWC model) opened dots represent the fitted by the best model data. To ease the interpretation, the results of the multiple regression functions, for which R_s was logarithmically transformed, were backtransformed to the original scale. Figure 5. Path diagrams showing the results of the Structural Equation Models (SEMs), represented by seasons. Arrows indicate causal relationships: positive and negative effects are indicated by solid and dashed arrows, respectively. Only the significant (p < 0.05) and marginally significant (p < 0.1) relationships were represented (see Table 3). The number given next to each arrow represents the Standardized Regression Weights (SRW) values given in Table 3. Path diagrams are represented in a plot where the X-axis represents the seasons (i.e., spring, summer, autumn, and winter) and the Y-axis represents the mean values of the soil

| respiration (R_s) flux for each season. Where, Grass, percentage of the soil surface covered by |
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| grass; T_s , soil temperature; R_s , soil respiration; $Slope$, micro-topography of the terrain within |
| the study stand; MeanD , mean distance from the European beech trees to the 81 measurement |
| points; $SWC,$ soil water content; $N_{trees},$ the count of all the surrounding European beech trees |
| around each of the 81 measurement points. |
| Figure 6. Conceptual framework illustrating how the observed spatio-temporal changes and |
| environmental controls of the soil respiration (R_s) respond to a seasonal shift that goes from |
| temperature-controlled (i.e., winter and autumn) to water-controlled (i.e., spring and summer) |
| processes. The X-axis represents the seasonal (i.e., winter, autumn, spring, and summer) soil |
| temperature (T_s) changes. The Y-axis represents the spatial variability of soil temperature (T_s) |
| and R _s represented as the standard deviation (SD) of the mean. The path diagrams, obtained |
| from the Structural Equation Models (SEMs; Figure 5), are also represented to show how the complexity of the controls of R_s increases along with the spatial variability of R_s . In the upper |
| part of the figure, the shift that goes from temperature-controlled (i.e., winter and autumn) to |
| water-controlled (i.e., spring and summer) processes over the spatial variability of R _s , is |
| indicated. The small figure panel included within the conceptual framework is represented by |
| Figure $4a$, with the red arrows indicating the seasonal temperature control of R_s (winter and |
| autumn) and the flattening of this control during warmer periods (spring and summer). Where, |
| N_{trees} , the count of all the surrounding European beech trees around each of the 81 measurement |
| points; SWC, soil water content; Grass, percentage of the soil surface covered by grass; Slope, |
| micro-topography of the terrain within the study stand; MeanD, mean distance from the |
| European beech trees to the 81 measurement points. |

| Spatial variability of soil respiration (R _s) and its controls are subjected to strong |
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| seasonality in an even-aged European beech (Fagus sylvatica L.) stand |
| |
| Running Title: Space-Time R _s variability in European beech stand |
| |
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| 31 | Keywords: | European | beech, | even-aged | stand, | micro-topography, | seasonality, | soi |
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microclimate, soil respiration, spatial variability

Abbreviations: $\mathbf{R_s}$, soil respiration; $\mathbf{T_s}$, soil temperature; \mathbf{SWC} , soil water content; \mathbf{BA} , basal area of the European beech trees surrounding the 81 measurement points; \mathbf{SOC} , soil organic carbon content; \mathbf{Litter} , thickness of the litter layer; \mathbf{Slope} , micro-topography of the terrain within the study stand; $\mathbf{N_{trees}}$, the count of all the surrounding European beech trees around each of the 81 measurement points; \mathbf{DBH} , average diameter at breast height (i.e., > 6 cm) of the European beech trees surrounding the 81 measurement points; \mathbf{MeanD} , mean distance from the European beech trees to the 81 measurement points; \mathbf{Grass} , percentage of the soil surface

covered by grass; Seedlings, percentage of the soil surface covered by tree seedlings.

Co Policy

Abstract

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Uncertainties arising from the so far poorly explained spatial variability of soil respiration (R_s) remain large. This is partly due to the limited understanding on how actually spatially variable R_s is, but also on how environmental controls determine R_s's spatial variability and how these controls vary in time (e.g., seasonally). Our study was designed to deepen into the complexity of R_s's spatial variability in a European beech even-aged stand, covering both phenologicaly and climatically contrasting periods (spring, summer, autumn, winter). Although we studied a relatively homogenous stand, we found a large spatial variability of R_s (coefficients of variation > 30%) characterized by strong seasonality. This large spatial variability of R_s suggests that even in relatively homogenous stands there is a large potential source of error when estimating R_s. This was also reflected by the sampling effort needed to obtain seasonal robust estimates of R_s, which may actually require a number of samples above that used in R_s studies. We further postulate that the effect of seasonality on the spatial variability and environmental controls of R_s was determined by the seasonal shifts of its microclimatic controls: during winter, low temperatures constrain plant and soil metabolic activities and hence reduce R_s variability (temperature-controlled processes), while during summer, water demand by vegetation and changes in water availability due to the micro-topography of the terrain (i.e., slope) increase R_s variability (water-controlled processes). This study provides novel information on the spatiotemporal variability of R_s and deepens into the seasonality of its environmental controls and the architecture of their causal-effect relationships controlling R_s's spatial variability. Our study further shows that improving current estimates of R_s at local and regional levels might be necessary in order to reduce uncertainties and improve CO₂ estimates at larger spatial scales.

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- The spatial variability of soil respiration (R_s) and its environmental controls vary seasonally
- Seasonal shifts from temperature- to water-controlled processes determine R_s's spatial
 variability
- Besides microclimate, slope and grass cover explain the spatio-temporal variability of $R_{\rm s}$
 - ❖ An intense sampling effort is needed to obtaining robust R_s estimates even in homogenous forests

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1. Introduction

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Soil respiration (R_s), i.e., the production and subsequent emission of carbon dioxide (CO₂) from the soil to the atmosphere, is one of the key processes contributing to the global terrestrial carbon (C) balance/budget. R_s is mainly produced by biological sources from the aerobic respiration of decomposers (i.e., heterotrophic respiration), as well as by plant roots and associated microorganisms living in the rhizosphere (i.e., autotrophic respiration) (Rodeghiero & Cescatti, 2008), but also by non-biological chemical oxidation reactions of C in organic matter, although at lower rates in this latter case (Raich & Schlesinger, 1992). Globally, the R_s emissions amount to a total of almost 80 PgC y⁻¹, being the second largest C flux after CO₂ uptake by plants (Raich & Tufekciogul, 2000), which means more than half of an ecosystem's total CO₂ emissions come from R_s (Barba et al., 2018; Curiel Yuste et al., 2005; Janssens et al., 2001). However, R_s is also probably the least well understood part of the C budget at global terrestrial ecosystems' level, based primarily on the fact that the large spatio-temporal variability that characterizes this large flux requires of a substantial monitoring effort at different scales and hence, of a large investment in instrumentation for its correct monitoring (Bond-Lamberty & Thomson, 2010). Therefore, and despite the large critical mass of studies performed to understand it (Bond-Lamberty & Thomson, 2010), our knowledge on the mechanisms controlling this large flux remains very limited (e.g., Barba et al., 2013; Curiel Yuste et al., 2019). Hence, there is still a need for studies designed to explore the spatiotemporal variability of R_s in order to be able to calibrate models and improve predictions of soil biological CO₂ emissions in a changing environment.

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The magnitude of R_s-related CO₂ emissions varies in time and space depending on multiple drivers. A critical mass of studies has been designed to understand how the temporal variability

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of R_s relates to different environmental factors such as soil temperature (T_s; e.g., Chen et al., 2014; Davidson et al., 1998; Epron et al., 2004a), soil water content (SWC; e.g., Davidson et al., 2000; Oishi et al., 2013; Poblador et al., 2017), wind (e.g., Sánchez-Cañete et al., 2016), or the photosynthetic activity of the plants (e.g., Bahn et al., 2009; Curiel Yuste et al., 2005; Davidson et al., 1998). Nevertheless, it is important to highlight the discrepancy between the large number of studies undertaken to understand the large, but predominantly explained variability in time (generally seasonal) of soil CO₂ fluxes (see for instance Bond-Lamberty & Thomson, 2010) and the relatively few studies undertaken to understand the enormous, but largely unexplained spatial variability of this very same flux. Several studies have proposed different factors that define local conditions as controls of the spatial variability of R_s at the mesoscale (scale of m). Most of these studies agree on the important role of variables such as: i). the spatial variability of soil moisture (Barba et al., 2013; Kosugi et al., 2007; Poblador et al., 2017); ii). the structure of the overstorey plant community (Barba et al., 2013; Epron et al., 2004b; Law et al., 2001; Saiz et al., 2006; Søe & Buchmann, 2005); iii). variables directly related to the structure of the aboveground plant community, such as leaf production (Oishi et al., 2013), root density or biomass (Knohl et al., 2008), microbial biomass, and litter thickness (Hanson et al., 1993); iv). the quantity and quality of soil organic matter (Rayment & Jarvis, 2000); or v). the C/N ratio and bulk density of the top soil (Khomik et al., 2006; Ngao et al., 2012; Saiz et al., 2006). Other topographical aspects, such as the slope and the position within the landscape, have been however less studied although their contribution to explain the spatial variability of R_s might also be critical (Arias-Navarro et al., 2017; Berryman et al., 2015; Brito et al., 2010; Hanson et al., 1993; Riveros-Iregui et al., 2012;). All studies, nevertheless, conclude that our capacity to predict the spatial variability of R_s and its environmental controls remains largely insufficient (e.g., Allaire et al., 2012).

The environmental controls of the spatial variability of R_s may also vary temporally, though, to the best of our knowledge, only few studies have been designed to deepen in this potential temporal axis of the spatial variation of R_s (Epron et al., 2004b; Khomik et al., 2006; Kosugi et al., 2007; Saiz et al., 2006; Shi et al., 2016; Søe & Buchmann, 2005). The complexity of the spatial variability of R_s can vary seasonally (Riveros-Iregui et al., 2012; Shi et al., 2016; Søe & Buchmann, 2005) specially because different environmental drivers may differently influence R_s depending on the season. For instance, the influence of water availability on the spatial patterns of R_s at the landscape-scale can exhibit a bidirectional behaviour, R_s being more sensitive to water availability during dry periods or in highly drained areas than during wetter periods or in low drainage areas (Riveros-Iregui et al., 2012). Likewise, the biomass and respiration of the autotrophic (roots and rhizosphere microorganism) and heterotrophic (microbial activity) components of R_s may vary in space and time depending on the phenological state of the vegetation and its nutrient and water demands (Barba et al., 2013; Søe & Buchmann, 2005). For this reason, understanding the drivers controlling the spatial variability of R_s at different temporal scales may help us to improve and modulate the sampling effort needed in order to obtain confident estimates of R_s. This also means that obtaining reliable integrative measures of R_s would require different sampling efforts throughout the year. It is, therefore, important to understand this seasonally-dependent complexity if we want to improve our knowledge on the sampling effort needed to get accurate and costly efficient estimates of R_s (Barba et al., 2013; Herbst et al., 2009; Rayment & Jarvis, 2000; Rodeghiero & Cescatti, 2008).

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We studied the spatio-temporal variability of soil respiration (R_s) in a 4.0 ha (i.e., 200 m x 200 m) European beech (*Fagus sylvatica* L.) even-aged stand. Specifically, we focused on understanding the potential seasonal (i.e., spring, summer, autumn, winter) variations of R_s and

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its environmental controls to: (1) determine the magnitude of their spatial variability and the sampling effort needed per each season (i.e., spring, summer, autumn, winter) to obtain robust average estimates of R_s; and (2) identify the main environmental controls and the architecture of their potential causal-effect relationships controlling the spatial variability of R_s along the seasons (i.e., during phenologically and climatically contrasted periods of the year). We hypothesized that, given the generally large influence of the aboveground plant distribution in explaining the spatial variability of R_s, the spatial variability of R_s will be low in our European beech even-aged study stand where trees are homogeneously distributed (H1). However, we also hypothesized that, along with other already well-studied and known factors, other factors (i.e., more spatially variable at stand level), such as the micro-topography of the terrain (i.e., slope) or the spatial distribution of the grass cover, will also play an important, indirect control over R_s due to their influence on the spatial variability of the soil water content (SWC) (H2). Finally, we also hypothesized that the predictive power of the different environmental controls of the spatial variability of R_s will vary throughout the year depending on the environmental constrains that act on R_s within a given season (e.g., soil temperature, T_s; or soil water content, SWC) (H3).

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2. Materials and Methods

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2.1. Study site and stand

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The study site (i.e., forest) is located in the central-southern part of Romania, in Mihaesti (Arges county; 45°05'11.8019"N, 25°03'58.0428"E), at an altitude of 570 m a.s.l. (Figure 1). This forest is largely dominated by European beech, although other tree species may also be found: hornbeam (*Carpinus betulus* L.), sessile oak (*Quercus petraea* (Matt.) Liebl.), or sweet cherry

(Prunus avium L.). The density within the whole forest located in Mihaesti is of 504 trees ha⁻¹, with a total volume of 502 m³ ha⁻¹, and a basal area of 33 m² ha⁻¹ (Mihaesti Forest Management Plan). Within this forest, we focused on a 4.0 ha European beech study stand (200 m x 200 m) (Figure 1). The European beech trees within the study stand are mainly adult and dominant (i.e., canopy level). According to the Mihaesti Forest Management Plan, most individuals within the study stand have an estimated age of ~ 85 years, which allows us to consider this study stand as being even-aged. The area where our study stand is located is characterized by a temperate continental climate, with a mean annual precipitation of ~ 875.21 mm, and a mean air temperature of ~ 6.31 °C, respectively (estimates calculated for the 1901 – 2019 period; CRU TS v.4; Harris et al., 2020). The mean annual precipitation for 2016 and 2017 (i.e., the years when our measurements were performed; see below) was of ~ 991.80 mm and ~ 959.60 mm, respectively. As for the mean annual air temperature, it was of ~ 7.59 °C in 2016 and of 7.54 °C in 2017 (CRU TS v.4; Harris et al., 2020). The soils are Eutric Cambisols (clay loam) covered with mull type humus, developed on a sandstone with marls parental material (Florea & Munteanu, 2012). The slope within the study stand is smooth and there are no important differences regarding the altitude between the upper part of the study stand and the lower part of the study stand (Figure 1, small panel). Mean pH values range from 4.8 (0-10 cm soil depth) to 5.2 (11-20 cm soil depth) (WTW pH330i; WTW GmbH, Weilheim, Germany).

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2.2. Field soil respiration (R_s) and microclimatic factors measurements

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The 4.0 ha selected study stand was divided into regular 25 m x 25 m squares (Figure 1, small panel). Soil respiration (R_s) measurements were then performed at each of the four corners of each of the 25 m x 25 m squares, resulting thus on a total of 81 measurement points. R_s measurements were all performed using a Portable Infrared Gas Analyzer (IRGA) connected to

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a soil respiration standard chamber (EGM-4 and SRC-1; PP Systems, Amesbury, MA, USA). The soil respiration chamber covered a soil surface area of 78 cm² and an enclosed volume of 1171 cm³. Since some studies have shown a clear correlation between insertion depth, the amount of cut roots, and the lost soil effluxes (Silvola et al., 1996; Wang et al., 2005), no collars were inserted in the soil (Arias-Navaro et al., 2017; Epron et al., 2004b; Hanson et al., 1993; Maestre & Cortina, 2003; Poblador et al., 2017). Instead, we followed a similar procedure to the one described by Epron et al., 2004b and we inserted the edge of the respiration chamber to a depth of 1 cm into the soil, including the litter layer. Nevertheless, this was done only after firstly removing the herbaceous layer in order to avoid potential confounding effects of the vegetation on R_s measurements. Furthermore, to avoid potential gas leaks due to the shallow insertion of the respiration chamber (1 cm into the soil) with respect to a relatively thick lowdensity litter layer (average 3.3 cm; Table 1), the respiration chamber was strongly pressed against the soil (i.e., with the help of one hand) over the whole time measurements were performed. Final R_s values were estimated for 120 seconds based on the linear increase of the CO₂ concentration within the soil respiration chamber (i.e., a closed dynamic system). Soil CO₂ efflux measurements were always performed between 9 a.m. and 5 p.m. Additionally, the CO₂ effluxes were never measured during rainy days. Specifically, in case of heavy rains (i.e., > 15 mm), field R_s measurements were postponed 36 h to avoid the "Birch effect" (Birch, 1958).

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Simultaneously to the field R_s measurements, microclimatic measurements (i.e., soil temperature and the volumetric soil water content) were also performed at the same 81 measurement points. Specifically, soil temperature (T_s) was measured at 5 cm soil depth using the STP-2 Soil Temperature Probe that was attached to the IRGA (PP Systems, Amesbury, MA, USA). As for the volumetric soil water content (SWC), this variable was measured at 20 cm soil depth using the TDR 300 soil moisture meter (Spectrum Technologies, Inc., Plainfield, IL,

USA). All field measurements (i.e., R_s, T_s, and SWC) spanned over a period of one complete year and thus over the four seasons: spring (May 2016), summer (August 2016), autumn (November 2016), and winter (February 2017). Within each of the 4 seasons and at each of the 81 measurement points, we performed 3 independent measurements for each of the 3 variables (i.e., R_s, T_s, and SWC) and then averaged their corresponding values. In order to systematically perform R_s, T_s, and SWC measurements at exactly the same locations within the study stand, we marked the 81 measurement points with wood sticks that were maintained in their positions over the whole study period. Due to the large number of measurement points (i.e., 81) and thus to the considerable field effort and logistics that were needed, R_s, T_s, and SWC measurements were always performed during 2 consecutive days during each season.

2.3. Forest structural and soil variables and the micro-topography of the terrain

At each of the 81 measurement points, soil samples were also collected to determine the soil organic carbon (SOC) content. All soil samples were collected in February 2017 after all seasonal measurements (i.e., R_s, T_s, and SWC) were finished. Soil sampling was performed using a metallic cylinder (5 cm diameter, and 20 cm depth) and consisted in extracting one soil core at each of the 81 measurement points. SOC of the upper 20 cm of the soil profile was determined through the dry combustion method using a CHNS organic elemental microanalyser (TruSpec Micro CHNS elemental analyser, LECO, New York, USA).

The thickness of the litter layer (hereinafter referred to as "litter" to simplify) was used as a proxy of litter biomass, which could not be measured due to logistics. The litter, at each of the 81 measurement points, was measured only once during the 2016 summer, two weeks before the R_s , T_s , and SWC measurements started. Although, we acknowledge the fact that it would

have been better to measure the litter layer over the year (i.e., seasons), this was not possible due to logistics. Instead, we assumed that the place where there was more accumulated litter (i.e., at some point) would be the same place where more litter usually falls and the opposite for the places where there was less accumulated litter. Accordingly, the litter depth would be basically stable over the year (i.e., seasons). Simultaneously to the litter measurements, the micro-topography of the terrain (hereinafter referred to as "slope" to simplify), at each of the 81 measurement points, was also measured.

In order to account for the impact of the surrounding vegetation on our field measurements (i.e., within a radius of 7 m around each of the 81 measurement points), we counted all the surrounding European beech trees (N_{trees}) and we measured their diameter at breast height (DBH; at standard 1.3 m above from the ground) and their distance to the 81 sampling points. The 7 m radius was established considering the average crown diameter of the European beech trees found within the 4.0 ha study stand (Mihaesti Forest Management Plan). The DBH of the trees was measured using a calliper (Haglöf, Sweden), only European beech trees with a DBH > 6 cm being finally considered for this study. The measured distances were used to calculate the mean distances (MeanD) from surrounding European beech trees to the 81 measurement points. In order to estimate the basal area (BA; m² ha⁻¹) of all European beech trees with a DBH > 6 cm, we calculated the sum of all their cross-sectional areas at breast height. Finally, within the same radius of 7 m around each of the 81 measurement points, we also estimated the percentage (%) of the soil surface covered by grass and the percentage (%) of the soil surface covered by all tree seedlings. These estimations were done visually and agreed between several observers for data consistency.

2.4. Statistical analyses

We used different statistics (i.e., mean, M; standard deviation, SD; and relative variability, RV) to estimate relative rates of spatial variability of forest structural and soil variables (i.e., BA, SOC, litter, N_{trees}, DBH, MeanD, % of grass, and % of seedlings) and of the micro-topography of the terrain (i.e., slope). We used the same statistics (i.e., mean, M; standard deviation, SD; and relative variability, RV) plus the absolute amplitude (A; defined as the difference between maximum and minimum values) to estimate relative and absolute rates of spatial variability of the microclimate (i.e., T_s and SWC) and soil respiration (R_s) variables. As most studies give the coefficient of variation (CV), we also calculated this statistic (i.e., expressed as a percentage) for the R_s variable alone and used it to compare our results with those published in previous studies. For the T_s, SWC, and R_s variables, all the above mentioned statistics were calculated considering both the four seasons separately and the four seasons combined (i.e., annual). The relative variability statistic (i.e., RV) was calculated following Lewontin 1966 and Webster 2001 as the standard deviation of the logarithms (i.e., log10 in our case) of measurements. Relative variability allows thus to compare variations between different groups of observations (Lewontin 1966; Webster 2001).

In order to calculate the minimum number of measurements (i.e., N) needed to obtain robust estimates of R_s for each season (i.e., spring, summer, autumn, and winter), we used the following power equation (Davidson et al., 2002):

$$N = [(t \times s)/(range/2)]^2 \qquad equation 1$$

where, t is represented by the critical value of the t-distribution (two-tailed test) for a given confidence level (99, 95, and 90%, respectively) and for 80 degrees of freedom; s is the standard

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deviation of all spatially independent R_s measurements per each season (i.e., spring, summer, autumn, and winter); and *range* is the width of the desired interval around the mean of the R_s measurements of each season (i.e., spring, summer, autumn, and winter) in which a smaller sample mean is expected to fall (i.e., error limit of 10%, 20%, and 30% of the R_s measurements mean per each season).

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We run geostatistical analyses (i.e., experimental (semi-) variograms and theoretical variograms) to determine the spatial autocorrelation of: i). the T_s, SWC, and R_s variables, separately for each season (i.e., spring, summer, autumn, and winter); ii). all forest structural and soil variables (i.e., basal area, BA; soil organic carbon content, SOC; thickness of the litter layer, litter; number of trees surrounding the 81 measurement points, N_{trees}; diameter at breast height, DBH; mean distance from surrounding European beech trees to the 81 measurement points, MeanD; percentages of grass and seedlings cover around the 81 measurement points); and iii). the micro-topography of the terrain (i.e., slope). Specifically, the experimental (semi-) variograms (i.e., binned) were performed using the "variog" function available from the "geoR" R package (Ribeiro et al., 2020) based on classical estimators. Further on, the theoretical variograms were performed using the "likfit" function available from the "geoR" R package (Ribeiro et al., 2020). These analyses were run considering the restricted maximum likelihood (REML) parameter estimation, different trends (i.e., the mean part of the model; constant, first order polynomial, and second order polynomial), and functions (i.e., models for the correlation function; matern, exponential, Gaussian, spherical, circular, cubic, wave, powered exponential, Cauchy, gneiting, and pure nugget). A total of 693 models were run. The selection of the best models for each of the analysed variables was based on AIC (Akaike Information Criteria). The selected models were then used to perform ordinary kriging using the "krige.conv" function available from the "geoR" R package (Ribeiro et al., 2020). The "image" function available from the "geoR" R package (Ribeiro et al., 2020) was finally used to visualize the spatial prediction based on the fixed covariance parameters. All selected models were validated using the "xvalid" function available from the "geoR" R package (Ribeiro et al., 2020).

To evaluate the microclimatic controls of the spatio-temporal variability of R_s , we run multiple regression functions using the "nls" function available from the "MASS" R package (Venables & Ripley, 2020). These functions were designed according to literature (i.e., Vicca et al., 2014). Briefly, different models were designed to represent the independent controls of soil temperature (T_s) and soil water content (SWC) on R_s , but also taking into account potential unimodal responses of R_s to both microclimatic factors. R_s data was log transformed prior to analyses as it did not meet the normality assumption.

$$\begin{split} \log(R_s) \sim a + b*T_s & \textit{equation 2} \\ \log(R_s) \sim a + b*SWC & \textit{equation 3} \\ \log(R_s) \sim a + b*T_s + c*T_s^2 & \textit{equation 4} \\ \log(R_s) \sim a + b*SWC + c*SWC^2 & \textit{equation 5} \\ \log(R_s) \sim a + b*T_s + c*SWC & \textit{equation 6} \\ \log(R_s) \sim a + b*T_s + c*SWC + d*T_s^2 & \textit{equation 7} \\ \log(R_s) \sim a + b*T_s + c*SWC + d*SWC^2 & \textit{equation 8} \\ \end{split}$$

Where, a, b, c, and d letters stand for coefficients of the multiple regression functions. The covariance and multicollinearity between T_s and SWC were examined prior to analyses using the Variance Inflation Factor (VIF). Since the VIF was lower than 2, both microclimatic variables could be used within the same model (Zuur et al., 2010). The selection of the best model was based on the AIC (Akaike Information Criterion). For these analyses, the T_s , SWC, and R_s

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variables were combined over the four seasons (i.e., spring, summer, autumn, and winter). All these analyses were conducted based on the assumption that the residuals of the models were normally distributed (p > 0.05) and independent. As the assumption of independence may be justified when data collection has been performed based on a probability sampling design (e.g., de Gruijter et al., 2006) and our data collection has been performed based on a grid sampling design, we acknowledge the fact that the spatial auto-correlation between data at nearby measurement points might impact some of the obtained results.

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To describe potential complex causal-effect relationships that might determine the spatiotemporal variability patterns of R_s, we run Structural Equation Models (SEMs). SEMs analyses allowed to test for the direct and indirect effects of all our measured in the field variables (i.e., microclimatic variables, forest structural and soil variables, and the micro-topography of the terrain) on seasonal (i.e., spring, summer, autumn, and winter) R_s. SEMs analyses were carried out using the "psem" function available from the "piecewiseSEM" R package (Lefcheck, 2016). To harmonize the results of the SEMs with those obtained from the multiple regression functions analyses, R_s was logarithmically transformed prior to analyses. Separated SEMs were built for each of the four seasons (i.e., spring, summer, autumn, and winter). All SEMs were designed based on hypotheses supported on simple univariate correlations between the different microclimatic (i.e., soil temperature, T_s; soil water content, SWC), micro-topography of the terrain (i.e., slope), soil (i.e., soil organic carbon content, SOC; and litter), and forest structural (i.e., basal area, BA; number of trees, N_{trees}; diameter at breast height, DBH; mean distance, MeanD; % of grass; and % of seedlings) variables (i.e., predictor variables). Furthermore, their potential complex causal-effect relationships, that might determine the spatio-temporal variability of R_s, were also considered. To test the goodness of fit of the SEMs, the Fisher's C statistic was calculated. The Fisher's C statistic follows a chi-squared distribution and tests if the model fits the data (p > 0.05) or not (p < 0.05). Several SEMs were run, the selection of the best one being based on the AIC (Akaike Information Criteria) (Lefcheck, 2016).

All statistical analyses were carried out in R (v. 4.0.0, R Core Team, 2020). Statistical relationships were considered significant at p < 0.05.

3. Results

3.1. The spatial variability of the forest structural and soil variables around the 81 soil respiration (R_s) measurement points

We found a relatively low spatial variability (i.e., expressed here as relative variability, RV; Table 1) for forest structural variables such as DBH (RV = 0.1) and MeanD (RV = 0.1) (Table 1, Figure S1). Similar results were obtained also for soil variables such as SOC (RV = 0.1) and litter (RV = 0.1) (Table 1, Figure S1). On the other hand, the spatial variability of the understory vegetation (i.e., % of grass and % of seedlings) showed the highest values and was higher relative to the spatial variability of the overstorey vegetation (i.e., BA and N_{trees}) (Table 1, Figure S1). High spatial variability values were found also for the slope (RV = 0.3) (Table 1, Figure S1).

| Statistics | BA | SOC | Litter | Slope | N _{trees} | DBH | MeanD | Grass | Seedlings |
|------------|-----------|-----------------------|--------|-------|-----------------------|------|-------|-------|-----------|
| Statistics | (m² ha-1) | (t ha ⁻¹) | (cm) | (%) | (N ha ⁻¹) | (cm) | (m) | (%) | (%) |
| M | 47.3 | 44.3 | 3.3 | 12.2 | 443 | 32.8 | 3.5 | 57.5 | 9.1 |
| SD | 16.9 | 12.9 | 0.7 | 8.3 | 200 | 6.3 | 0.8 | 31.1 | 8.6 |
| RV | 0.2 | 0.1 | 0.1 | 0.3 | 0.2 | 0.1 | 0.1 | 0.4 | 0.3 |

Table 1. Mean (M), standard deviation (SD), and relative variability (RV; following Lewontin 1966 and Webster 2001) values of forest structural, soil, and the micro-topography of the terrain variables. *Where*, **BA**, basal area of the European beech trees surrounding the 81 measurement points; **SOC**, soil organic carbon content; **Litter**, thickness of the litter layer; **Slope**, microtopography of the terrain within the study stand; **N**_{trees}, the count of all the surrounding European beech trees around each of the 81 measurement points; **DBH**, average diameter at breast height (i.e., > 6 cm) of the European beech trees surrounding the 81 measurement points; **MeanD**, mean distance from the European beech trees to the 81 measurement points; **Grass**, percentage of the soil surface covered by grass; **Seedlings**, percentage of the soil surface covered by tree seedlings.

3.2. Soil respiration (R_s) sampling effort needed per season

The calculation of the minimum number of measurements needed to obtain robust estimates of R_s for each season (i.e., spring, summer, autumn, and winter) showed how sensitive *equation 1* was to both the error limit (i.e., 10%, 20%, and 30%) and the confidence interval (90%, 95%, and 99%) (Table S1). Accordinly, the calculated sampling effort varied within and among seasons depending on the error limit and the confidence interval. Specifically, the sampling effort varied more strongly within seasons than among them (Table S1). As for the sampling effort needed from one season to another to obtain robust estimates of R_s , differences were not so strong especially between summer, autumn, and winter (Table S1). Spring was the season when less minimum R_s measurements seem to be needed to obtain robust estimates of R_s (Table S1) no matter the confidence level and the error limit (Table S1). On the other hand, winter was found to be the season when more sampling effort seems to be needed to obtain robust estimates of R_s no matter the confidence level and the error limit (Table S1). The sampling effort for the

| summer season seems to be quite close to the winter one, while the sampling effort for the |
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| autumn season was found to be slightly lower (i.e., in between the spring and summer), no |
| matter the confidence level and the error limit (Table S1). |
| |

3.3. Seasonal and spatial variability of soil respiration (R_s) , soil temperature (T_s) , and soil water content (SWC)

As expected, both microclimatic variables (i.e., T_s and SWC) experienced very different seasonal patterns during the study period (Table 2, Figure 2). On one hand, T_s experienced large seasonal changes, peaking during summer and reaching its minimums during winter (Table 2, Figure 2a). On the other hand, values of SWC experienced less seasonality, reaching its minimums during summer, but being very stable and similar for the rest of the year (Table 2, Figure 2b). Seasonality of R_s followed a pattern similar to that observed for T_s , peaking during both spring and summer, and reaching its minimums during winter (Table 2, Figure 2c).

The spatial variability of the microclimatic variables was also markedly different (Table 2, Figure 3). Specifically, the spatial variability of T_s was generally smaller (RV values ranging from 0.01 to 0.10), within the range of 3.1 to 5.3 °C of amplitude (Table 2, Figure 3A, D, G, J), than the spatial variability of SWC (RV values ranging from 0.10 to 0.14), within the range of 31.9 to 54.6 % vol. of amplitude (Table 2, Figure 3B, E, H, K). R_s was the variable with the highest spatial variability (RV values ranging from 0.13 to 0.17), within the range of 2.5 to 11.9 μ mol CO_2 m⁻² s⁻¹ of amplitude (Table 2, Figure 3C, F, I, L) and coefficients of variation above 30% (Table 2). As concerning the standard deviation of the mean values, both T_s and SWC showed seasonal changes, summer being the season that showed the lowest values (Table 2). Standard deviation values for R_s also showed seasonal changes, being higher in spring and

summer and smaller in winter (Table 2), a pattern that followed the increase of the R_s rates (Figure 2).

| | | | Ts | | | SV | VC | | | | R_s | | |
|--------|------|-----|------|------|----------|------|------|------|---|-----|-------|------|------|
| Period | (°C) | | | | (% vol.) | | | | (μmol CO ₂ m ⁻² s ⁻¹) | | | | |
| | M | SD | RV | A | M | SD | RV | A | M | SD | RV | CV | A |
| Spring | 14.4 | 0.9 | 0.03 | 4.6 | 33.0 | 10.2 | 0.14 | 54.6 | 6.3 | 2.0 | 0.13 | 31.1 | 8.7 |
| Summer | 17.4 | 0.6 | 0.01 | 3.1 | 24.9 | 6.6 | 0.12 | 31.9 | 6.4 | 2.4 | 0.17 | 37.6 | 11.9 |
| Autumn | 7.5 | 0.7 | 0.04 | 3.9 | 29.6 | 8.1 | 0.11 | 42.7 | 2.8 | 1.0 | 0.15 | 34.9 | 4.8 |
| Winter | 3.8 | 0.9 | 0.10 | 5.3 | 33.9 | 8.0 | 0.10 | 41.9 | 1.3 | 0.5 | 0.16 | 36.4 | 2.5 |
| Annual | 10.8 | 5.5 | 0.27 | 17.7 | 30.4 | 9.0 | 0.13 | 54.6 | 4.2 | 2.7 | 0.33 | 65.1 | 13.6 |

Table 2. Mean (M), standard deviation (SD), relative variability (RV; following Lewontin 1966 and Webster 2001), Coefficient of Variation (CV; only for the R_s variable) expressed as a percentage, and absolute amplitude (A; defined as the difference between maximum and minimum values) values of soil microclimate (i.e., T_s, soil temperature; and SWC, soil water content) and soil respiration (R_s) measurements. The above mentioned statistics have been calculated both at the seasonal level (i.e., spring, summer, autumn, and winter) and over the four seasons combined (i.e., annual).

3.4. The spatio-temporal variability of soil respiration (Rs)

The model that best explained the microclimatic controls over the spatio-temporal variability of R_s was the model that considered unimodal responses of R_s to T_s and a negative linear response of R_s to SWC (i.e., equation 7) (Table S2, Figure 4). Specifically, the unimodal effect of T_s on R_s translated into a seasonal sensitivity of R_s to T_s : i.e., R_s response to T_s was stronger at lower T_s values (i.e., corresponding to autumn and winter seasons) than at higher T_s values

(i.e., corresponding to spring and summer seasons) (Figure 4a) when R_s reached its peak (Table 2) and its response to the seasonal changes of T_s was low. The overall effect of SWC over the spatio-temporal variability of R_s was negative, this response being evidenced by the negative slope of this relationship (Figure 4b). This negative effect was further on highlighted by the overall relationship that was found between SWC and R_s at the spatial scale (i.e., results of the SEM analyses), where higher values of SWC were generally associated with low R_s rates (see Figures 4b and 5). Hence, the SWC effect on R_s was mainly driven by the spatially net negative effect of SWC on R_s .

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Both multiple regression functions and SEMs agreed that the microclimatic variables (T_s and SWC; Figures 4 and 5) were, overall, the most important factors controlling the variability of R_s in our 4.0 ha study stand. Nevertheless, SEMs further showed how the architecture of the potential causal-effect relationships controlling R_s's spatial variability increased in complexity (Figure 5) during spring and summer, when the spatial variability of R_s was maximal (see SD and A in Table 2 and Figure 2C). Specifically, SEMs highlighted how during the summer season, when the spatial variability of R_s was the highest (Table 2), the number of variables ultimately involved in controlling the spatio-temporal variability of R_s were also high compared to, e.g. winter or autumn (Figure 5). Also, the predictive capacity of the spatio-temporal variability of R_s varied seasonally (Table 3), the coefficient of determination (R²) ranging from 0.10 in winter to 0.29 in autumn. SEMs showed how both the forest structural (N_{trees} , MeanD, and % of grass cover) and the micro-topography of the terrain (i.e., slope) variables may strongly influence, directly and indirectly the spatio-temporal variability of R_s (Figure 5). Specifically, in spring and summer, the % of grass cover was negatively associated with T_s and SWC, which on their turn exerted a further positive and negative, respectively, influence on the spatio-temporal variability of R_s (Figure 5). N_{trees} instead, always showed a direct and positive

effect on the spatio-temporal variability of R_s during the coldest seasons (i.e., autumn and winter; Figure 5). The slope (i.e., the micro-topography of the terrain within the study stand) was negatively related with MeanD in spring and summer, this relationship being especially important during summer when MeanD exerted some control (i.e., positive relationship) over R_s (Figure 5).



| Season | Response | Predictor | Estimate | SE | df | p-value | SRW | Response | R² | n | Fisher's C | df | p-value |
|--------|---------------------------|--------------------|----------|--------|----|---------|-------|----------------|------|----|------------|----|---------|
| | $\mathbf{R}_{\mathbf{s}}$ | T_s | 0.0583 | 0.0154 | 78 | 0.0003 | 0.39 | R _s | 0.17 | 81 | 10.193 | 10 | 0.424 |
| S | R_s | Slope | -0.0031 | 0.0017 | 78 | 0.0672 | -0.19 | MeanD | 0.05 | | | | |
| Spring | MeanD | Slope | -0.0213 | 0.0103 | 79 | 0.0413 | -0.23 | T _s | 0.07 | | | | |
| | T_s | Grass | -0.0076 | 0.0031 | 79 | 0.0170 | -0.26 | | | | | | |
| | $\mathbf{R}_{\mathbf{s}}$ | SWC | -0.0086 | 0.0028 | 76 | 0.0030 | -0.32 | R _s | 0.24 | 80 | 12.3 | 12 | 0.422 |
| | R_s | T_s | -0.0836 | 0.0320 | 76 | 0.0109 | -0.27 | SWC | 0.11 | | | | |
| C | R_s | MeanD | 0.0471 | 0.0227 | 76 | 0.0412 | 0.22 | MeanD | 0.05 | | | | |
| Summer | SWC | Slope | -0.1890 | 0.0829 | 77 | 0.0253 | -0.25 | | | | | | |
| | SWC | Grass | -0.0571 | 0.0226 | 77 | 0.0134 | -0.28 | | | | | | |
| | MeanD | Slope | -0.0213 | 0.0103 | 79 | 0.0413 | -0.23 | | | | | | |
| | $\mathbf{R}_{\mathbf{s}}$ | T_s | 0.0504 | 0.0233 | 77 | 0.0331 | 0.22 | R _s | 0.29 | 81 | 1.581 | 2 | 0.454 |
| | R_s | SWC | -0.0076 | 0.0018 | 77 | 0.0001 | -0.40 | $T_{\rm s}$ | 0.07 | | | | |
| Autumn | R_s | N _{trees} | 0.0002 | 0.0001 | 77 | 0.0486 | 0.20 | | | | | | |
| | T _s | N _{trees} | 0.0008 | 0.0004 | 79 | 0.0190 | 0.26 | | | | | | |
| Winter | R_s | N _{trees} | 0.0003 | 0.0001 | 79 | 0.0045 | 0.31 | $R_{\rm s}$ | 0.10 | 81 | 0 | 0 | 1 |

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Table 3. Statistics of the Structural Equation Models (SEMs) analyses showing causal-effect relationships that determine the spatio-temporal variability of soil respiration (R_s). Only significant (p < 0.05) and marginally significant (p < 0.1) causal relationships are given. Where, T_s , soil temperature; SWC, soil water content; Slope, micro-topography of the terrain within the study stand; N_{trees} , the count of all surrounding European beech trees around each of the 81 measurement points; MeanD, mean distance from the European beech trees to the 81 measurement points; Grass, percentage of the soil surface covered by grass; SE, standard error; df, degrees of freedom; SRW, Standardized Regression Weights; R², the coefficient of determination; \mathbf{n} , sampling size; **Fisher's C statistic**, follows a chi-squared distribution and tests if the model fits the data (p > 0.05) or not (p < 0.05). The right hand part of the table shows the statistics of the best models representing the spatio-temporal variability of R_s during each of the four different seasons (i.e., spring, summer, autumn, and winter). r).

4. Discussion

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We here report high soil respiration (R_s) coefficient of variation values (i.e., CV, ranging from 31.1 in spring to 37.6 in summer; see Table 2) in an even-aged European beech study stand located in the central-southern part of Romania (Mihaesti, Arges county). These values are higher or comparable to other CV values mentioned in previous studies (e.g., Barba et al., 2013; Epron et al., 2004b; Kosugi et al., 2007; Ngao et al., 2012; Shi et al., 2016; Stoyan et al., 2000), although caution should be taken when comparing CV values among studies as they might also vary depending on the considered spatial scales (e.g., Darenova & Čater, 2020; Ngao et al., 2012). Nevertheless, independent of this consideration, the high CV values we obtained here refute our first hypotheses (H1). In fact, the magnitude of the spatial variability of R_s during the warmest seasons (i.e., spring and summer) was comparable to the overall annual variability of R_s (see SD and A values in Table 2), which reinforces the idea of the large, though generally neglected, impact of R_s's spatial variability on estimates of soil CO₂ effluxes, even in homogenous ecosystems such as the European beech even-aged study stand that we considered here. The calculated large sampling effort needed to obtain robust estimates of R_s for any given season (being even larger in winter, see Table S1) further highlights the importance of the spatial variability of R_s as a potential source of uncertainty on local and global CO₂ estimates and that should be taken into account. This is of upmost importance especially now, when the number of studies dedicated to scale up CO₂ observations from local to global levels is growing. Accordingly, our study suggests that obtaining robust estimates of R_s at the local level may require of more intense spatial sampling efforts, than those generally carried out for logistical reasons, in order to address and diagnose uncertainties on CO₂ estimates at the global level (e.g., Jian et al., 2018; Warner et al., 2019).

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The large spatio-temporal variability of R_s was strongly and directly determined by soil microclimatic conditions (T_s and SWC; Figures 4 and 5). Nevertheless, as hypothesized (H2), less considered variables such as the forest structural ones (i.e., grass, MeanD, N_{trees}) or the micro-topography of the terrain (i.e., slope), proved to have a determinant, direct or indirect, effect on the observed spatio-temporal variability of R_s. In the case of the slope and the grass cover variables, they both showed further tight relationships with soil microclimatic conditions (i.e., T_s and SWC). These relationships were found to be significant in spring and summer (Figure 5), the two seasons when R_s values peaked (Table 2). Specifically, grass cover modulated the T_s variable in spring, with an indirect effect over R_s. In summer instead, when SWC usually registers low values and the competition for water and nutrients between the heterotrophic communities and the vegetation is high (Villegas et al., 2010), grass cover modulated the SWC availability, with an indirect effect over R_s. Our results highlight thus the importance of seldom considered variables, such as the micro-topography of the terrain (e.g., Arias-Navarro et al., 2017) and the vegetation (e.g., Søe & Buchmann, 2005), in R_s studies, as they may actually substantially impact, either directly or indirectly, the spatio-temporal variability of R_s. Instead, in our European beech even-aged study stand, we found no significant effects of variables generally well associated with the spatial variability of R_s, such as the litter (e.g., Epron et al., 2004b; Katayama et al., 2009; Saiz et al., 2006) or the soil organic carbon content (e.g., Søe & Buchmann, 2005). Although the fact that litter thickness was only measured once (i.e., during the 2016 summer; cf. 2.3. section), and thus may have generated a certain source of noise in our models (since the litter generally accumulates in autumn in deciduous-dominated forests), we assumed that our summer measurements contain very valuable information on the long-term spatial patterns of litter accumulation on the soil, and therefore, valuable information on where litter can have a greater impact on the spatial variability of soil processes in the long term. We are further aware of the limitations of our study, since other variables, such as soil compaction (Schwen et al., 2015) or the spatial distribution of the root biomass (Søe & Buchmann, 2005), that have not been measured, may have also helped to explain the observed spatial patterns of R_s. However, we expect that effects of the spatial variability of soil compaction on, e.g. water infiltration or CO₂ diffusivity (e.g., Schwen et al., 2015), will not be as high as in more intensively used stands since our study stand has not undergone any forestry intervention during the last 85 years and has no livestock load (according to the Mihaesti Forest Management Plan). On the other hand, and given the logistical inability to obtain estimates of the spatial distribution of root biomass, our exhaustive characterization of the distribution of trees and understorey (e.g., grass and seedlings) around the 81 measurement points emerged as a good proxy highly associated with the distribution of roots, assuming that proximity to vegetation is closely associated with root density in the soil (Søe & Buchmann, 2005).

Our results further emphasized the importance of understanding the temporal (i.e., seasonal) changes in the magnitude and controls of spatial variability of R_s . This variability could be especially important in temperate areas where microclimatic conditions (i.e., T_s and SWC), soil CO_2 effluxes, and vegetation activity may vary dramatically throughout the year (e.g., Curiel Yuste et al., 2005). Indeed, and also as hypothesized (H3), our results indicated that, along with the observed seasonal variability in the magnitude of R_s , the variables that control the spatial variability of R_s were also subjected to strong seasonality. The architecture of the causal-effect relationships controlling R_s 's spatial variability varied between the four seasons and showed an increased complexity during spring and summer, while in autumn and especially in winter these relationships were much simpler (Figure 5). In line with these findings, spring and summer were also the seasons when the highest R_s values were registered, as expected for temperate ecosystems (e.g., Knohl et al., 2008; Saiz et al., 2006; Shi et al., 2016; Søe & Buchmann, 2005).

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These high R_s values coincided thus with the warmest temperatures of the year, with the peak in plant and soil biological activity, and with the highest variability of R_s in absolute terms. These results are of upmost importance as they highlight the fact that in order to obtain robust estimates of R_s -CO₂ derived emissions and to have a deeper understanding on the R_s variability, both spatial and temporal R_s controlling processes need to be taken into account.

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Based on our results, we postulate that in this even-aged European beech study stand, the observed spatio-temporal changes and controls of the R_s respond to a seasonal shift that goes from temperature-controlled (i.e., winter and autumn) to water-controlled (i.e., spring and summer) processes. Figure 6 shows a conceptual framework, based on our results, that illustrates this shift in time. During cold periods, when the seasonal variability of R_s was, as expected in a temperature forest (e.g., Curiel Yuste et al. 2003), strongly limited by T_s (Figure 4a), the spatial variability in R_s followed the low spatial variability of T_s (represented as standard deviation of the mean R_s or T_s in Figure 6), resulting in low spatial variability of R_s (Table 2, Figure 6). Most factors had insignificant effects over the spatial variability of R_s during the winter season, when only N_{trees} (i.e., the number of European beech trees surrounding the 81 measurement points) showed a positive relationship with R_s (Figure 5). These results might be related with a larger autotrophic respiration contribution to R_s during winter when European beech trees are able to maintain part of their fine root biomass alive (e.g., Büttner and Leuschner, 1994; Zwetsloot et al., 2019). A similar result was found for the autumn season, when N_{trees} also showed a positive relationship with R_s (Figure 5). During warm periods (i.e., summer), when the soil metabolic activity is at its peak (reflected in higher rates of R_s; Table 2, Figure 2), the increase in temperature and vegetation activity increases the demand for SWC (evapotranspiration), which then becomes a limiting factor for R_s. Although our SWC measurements were too deep (i.e., 20 cm soil depth) to capture this increasing water control

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(i.e., our model could not capture a positive effect of SWC on R_s; see Figure 4b), this seasonal water limitation of R_s was evidenced by the low sensitivity to temperature that R_s experienced during the warmer periods (Figure 4a), which is the shape typically observed when R_s responds to a shift that goes from temperature-controlled to water-controlled processes (e.g., Curiel-Yuste et al., 2003, 2005; Davidson & Janssens, 2006). This shift towards R_s's spatial variability being water-controlled resulted in an increase in R_s variability, which subsequently become more spatially variable than temperature (reflected in increased standard deviation values of R_s with respect to T_s; see Figure 6). The shift towards a water-limited R_s system that generated spatial variability of R_s (Figure 6) also increased the complexity of R_s controls (Figures 5 and 6). This is because the increase in vegetation activity triggered a higher competition for water, as evidenced by, e.g. the strong negative influence of tree proximity (i.e., positive effect of MeanD in summer on R_s; Figure 5) or the strong negative effect of the grass cover over SWC during summer (Figure 5). Hence, the evaporative demand of the vegetation (i.e., MeanD and grass cover) exerted direct and indirect controls over the spatial variability of R_s during dry, warm, and phenologically active periods (i.e., spring and summer), contributing to an increase in the spatial variability of R_s. The slope (i.e., the micro-topography of the terrain) was another variable that contributed, directly and indirectly, to the increase in the spatial variability of R_s during warmer periods characterized by higher water demand (Figure 5). Slope may have large impacts over water availability and water balances by creating spatial variability in e.g. the incidence of solar radiation at the floor level and water run-off (Berryman et al., 2015; Riveros-Iregui et al., 2012), with further consequences on soil CO₂ effluxes, even across short distances (Arias-Navarro et al., 2017). It is likely that, at our study stand, the spatial distribution of the slopes captured the spatial variability of SWC during drier periods (i.e., summer) better than our own SWC measurements taken at 20 cm depth. This is because during periods of high water demand (i.e., summer), SWC decreases very fast in the uppermost layer of the soil, where most

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of both autotrophic and heterotrophic activities concentrate (Curiel Yuste et al., 2003, 2005), whereas at 20 cm depth SWC remains above the volumetric content thresholds at which SWC limits Rs, as stated by the fact that no positive relationship was found between SWC and $R_{\rm s}$ (Figures 4b and 5).

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5. Conclusions

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We here highlight the fact that the spatial variability of R_s proves to be high even in a relatively homogenous even-aged European beech study stand of 4.0 ha. Accordingly, our estimates regarding the sampling effort needed to obtain robust estimates of R_s further suggest that most studies to date might have probably underestimated the sampling effort needed to obtain accurate spatial estimates of R_s throughout the year. Our study further shows that the spatial variability of R_s, varied significantly throughout the year, peaking in spring and summer and being low in winter, coinciding thus with the seasonal variability in the absolute magnitude of R_s. We here postulate that in this European beech-dominated even-aged study stand, the observed large seasonal changes in the magnitude and controls of the spatial variability of R_s respond to a seasonal shift that goes from temperature-controlled (i.e., winter and autumn) to water-controlled (i.e., spring and summer) processes. This is because when temperatures and water demands are high, the evaporative demand of both the overstorey but also the understorey vegetation, as well as the micro-topography of the terrain (i.e., slope), generate spatial complexity in soil R_s. During winter, temperature limits processes and prevents most other factors from spatially influencing R_s. In conclusion, obtaining robust, accurate estimates of R_sderived CO₂ effluxes, may profit from: (1) a deeper understanding of how the spatial patterns of R_s varies across seasons, e.g., understanding when processes shift from being controlled by temperature (i.e., winter and autumn) to being controlled by water (i.e., spring and summer);

and (2) a deeper understanding on how, when, and where, factors such as the micro-topography of the terrain or the plant-plant and the plant-soil competition for water may contribute to this spatial variability of R_s. In line with our findings, it would be interesting to test in future at which extent this observed trends apply to other types of ecosystems or if they may also be extrapolated to latitudinal and/or altitudinal gradients, i.e., whether R_s's spatial complexity may increase considering gradients that go from temperature-limited (e.g., temperate) to water-limited (e.g., arid and semi-arid) systems, or from topographically simple (e.g., valleys) to topographically more complex (e.g., mountains) systems.

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Figure 1. Map indicating the location of Romania within Europe and the location of the study site in the central-southern part of Romania (i.e., Mihaesti, Arges county). The small panel shows the sampling design: a 4.0 ha (i.e., 200 m x 200 m) even-aged European beech stand divided into 25 m x 25 m squares. The scale that appears on the right size of the small panel indicates the altitude (m a.s.l.) gradient within the study stand. Figure 2. Seasonal (i.e., spring, summer, autumn, and winter) patterns of: a) soil temperature (T_s) ; b) soil water content (SWC); and c) soil respiration (R_s) . Figure 3. Spatial prediction based on the fixed covariance parameters generated by performing geostatistical analyses on the seasonal (i.e., spring, summer, autumn, and winter) spatial distribution of soil temperature (T_s), soil water content (SWC), and soil respiration (R_s). **Figure 4**. Representation of the best model (Table S2) that explained the microclimatic controls (soil temperature, T_s; and soil water content, SWC) over the spatio-temporal variability of soil respiration (R_s): a) R_s response to T_s; and b) R_s response to SWC. Black opened dots represent the row data, while red (i.e., R_s response to T_s model) and blue (i.e., R_s response to SWC model) opened dots represent the fitted by the best model data. To ease the interpretation, the results of the multiple regression functions, for which R_s was logarithmically transformed, were backtransformed to the original scale. Figure 5. Path diagrams showing the results of the Structural Equation Models (SEMs), represented by seasons. Arrows indicate causal relationships: positive and negative effects are indicated by solid and dashed arrows, respectively. Only the significant (p < 0.05) and marginally significant (p < 0.1) relationships were represented (see Table 3). The number given next to each arrow represents the Standardized Regression Weights (SRW) values given in Table 3. Path diagrams are represented in a plot where the X-axis represents the seasons (i.e., spring, summer, autumn, and winter) and the Y-axis represents the mean values of the soil

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respiration (R_s) flux for each season. Where, Grass, percentage of the soil surface covered by grass; T_s , soil temperature; R_s , soil respiration; Slope, micro-topography of the terrain within the study stand; MeanD, mean distance from the European beech trees to the 81 measurement points; SWC, soil water content; N_{trees} , the count of all the surrounding European beech trees around each of the 81 measurement points. Figure 6. Conceptual framework illustrating how the observed spatio-temporal changes and environmental controls of the soil respiration (R_s) respond to a seasonal shift that goes from temperature-controlled (i.e., winter and autumn) to water-controlled (i.e., spring and summer) processes. The X-axis represents the seasonal (i.e., winter, autumn, spring, and summer) soil temperature (T_s) changes. The Y-axis represents the spatial variability of soil temperature (T_s) and R_s represented as the standard deviation (SD) of the mean. The path diagrams, obtained from the Structural Equation Models (SEMs; Figure 5), are also represented to show how the complexity of the controls of R_s increases along with the spatial variability of R_s. In the upper part of the figure, the shift that goes from temperature-controlled (i.e., winter and autumn) to water-controlled (i.e., spring and summer) processes over the spatial variability of R_s, is indicated. The small figure panel included within the conceptual framework is represented by Figure 4a, with the red arrows indicating the seasonal temperature control of R_s (winter and autumn) and the flattening of this control during warmer periods (spring and summer). Where, N_{trees}, the count of all the surrounding European beech trees around each of the 81 measurement points; SWC, soil water content; Grass, percentage of the soil surface covered by grass; Slope, micro-topography of the terrain within the study stand; MeanD, mean distance from the European beech trees to the 81 measurement points.

Supplementary material

| | Minimum soil respiration (R _s) | | | | | | | | | |
|--------|--|--------------------|-----|-----|--|--|--|--|--|--|
| Season | measurements number (i.e., N) | | | | | | | | | |
| | Error limit | Confidence interva | | | | | | | | |
| | | 90% | 95% | 99% | | | | | | |
| | ±10% | 112 | 160 | 281 | | | | | | |
| Spring | ±20% | 28 | 40 | 70 | | | | | | |
| | ±30% | 12 | 18 | 31 | | | | | | |
| | ±10% | 156 | 223 | 392 | | | | | | |
| Summer | ±20% | 39 | 56 | 98 | | | | | | |
| | ±30% | 17 | 25 | 44 | | | | | | |
| | ±10% | 141 | 202 | 355 | | | | | | |
| Autumn | ±20% | 35 | 51 | 89 | | | | | | |
| | ±30% | 16 | 22 | 39 | | | | | | |
| | ±10% | 164 | 234 | 412 | | | | | | |
| Winter | ±20% | 41 | 59 | 103 | | | | | | |
| | ±30% | 18 | 26 | 46 | | | | | | |
| | | | | | | | | | | |

Table S1. Minimum number of measurements (i.e., N) needed to obtain robust estimates of soil respiration (R_s) for each season (i.e., spring, summer, autumn, and winter) according to *equation 1* (cf. 2.4. Statistical analyses). Results are given for 90%, 95%, and 99% confidence intervals. The *range* is the width of the desired interval around the mean of the R_s measurements of each season (i.e., spring, summer, autumn, and winter) in which a smaller sample mean is expected to fall (i.e., error limit of 10%, 20%, and 30% of the R_s measurements mean per each season).

| Model | | a | b | c | d | AIC | R ² |
|---|---|-----------|------------|------------|------------|--------|----------------|
| $\log(R_s) \sim a + b * T_s$ | | 0.6002*** | 0.0843*** | | | 125.30 | 0.85 |
| $log(R_s) \sim a + b*SWC$ | | 2.0141*** | -0.0166*** | | | 505.55 | 0.27 |
| $log(R_s) \sim a + b T_s + c T_s^2$ | | 0.1483* | 0.2002*** | -0.0055*** | | 83.45 | 0.87 |
| $log(R_s) \sim a + b*SWC + c*SWC^2$ | | 2.9595*** | -0.0772*** | 0.0009*** | | 491.43 | 0.35 |
| $\log(R_s) \sim a + b^*T_s + c^*SWC$ | 1 | 0.7544*** | 0.0825*** | -0.0044* | | 121.50 | 0.85 |
| $log(R_s) \sim a + b*T_s + c*SWC + d*T_s^2$ | | 0.3144*** | 0.2008*** | -0.0051** | -0.0056*** | 76.76 | 0.87 |
| $log(R_s) \sim a + b*T_s + c*SWC + d*SWC^2$ | 4 | 1.0960*** | 0.0812*** | -0.0252** | 0.0003* | 117.65 | 0.85 |

Table S2. Results of multiple regression functions used to explain the spatio-temporal variability of soil respiration (R_s). The table shows the models that were considered (i.e., following Vicca et al., 2014), the estimated coefficients of the multiple regression (a, b, c, and d), the AIC (Akaike Information Criterion) values, and the coefficient of determination (R^2) values. The best model, based on AIC, is marked in bold. *Where*, superscript asterisks of the estimated coefficients of the multiple regression (i.e., a, b, c, and d) stand for: p-values < 0.05 (*), p-values < 0.01 (***), and p-values < 0.001 (***), respectively.

Figure caption:

Figure S1. Spatial prediction based on the fixed covariance parameters generated by performing geostatistical analyses on the seasonal spatial distribution of: **BA**, basal area of the European beech trees surrounding the 81 measurement points (panel A); SOC, soil organic carbon content (panel B); Litter, thickness of the litter layer (panel C); Slope, micro-topography of the terrain within the study stand (panel D); N_{trees} , the count of all the surrounding European beech trees around each of the 81 measurement points (panel E); **DBH**, average diameter at breast height (i.e., > 6 cm) of the European beech trees surrounding the 81 measurement points (panel F); MeanD, mean distance from the European beech trees to the 81 measurement points (panel G); Grass, percentage of the soil surface covered by grass (panel H); Seedlings, percentage of the soil surface covered by tree seedlings (panel I).