Archival Report

High Polygenic Risk Scores Are Associated With Early Age of Onset of Alcohol Use Disorder in Adolescents and Young Adults at Risk

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ABSTRACT

BACKGROUND: Genome-wide association studies have been conducted in alcohol use disorder (AUD), and they permit the use of polygenic risk scores (PRSs), in combination with clinical variables, to predict the onset of AUD in vulnerable populations.

METHODS: A total of 2794 adolescent/young adult subjects from the Collaborative Study on the Genetics of Alcoholism were followed, with clinical assessments every 2 years. Subjects were genotyped using a genomewide chip. Separate PRS analyses were performed for subjects of European ancestry and African ancestry. Age of onset of DSM-5 AUD was evaluated using the Cox proportional hazard model. Predictive power was assessed using receiver operating characteristic curves and by analysis of the distribution of PRS.

RESULTS: European ancestry subjects with higher than median PRSs were at greater risk for onset of AUD than subjects with lower than median PRSs ($p = 3 \times 10^{-7}$). Area under the curve for the receiver operating characteristic analysis peaked at 0.88 to 0.95 using PRS plus sex, family history, comorbid disorders, age at first drink, and peer drinking; predictive power was primarily driven by clinical variables. In this high-risk sample, European ancestry subjects with a PRS score in the highest quartile showed a 72% risk for developing AUD and a 35% risk of developing severe AUD (compared with risks of 54% and 16%, respectively, in the lowest quartile).

CONCLUSIONS: Predictive power for PRSs in the extremes of the distribution suggests that these may have future clinical utility. Uncertainties in interpretation at the individual level still preclude current application.

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Alcohol use disorder (AUD) has its peak onset in late adolescence and early adulthood [ages 18–29 years (1)]. Family history of DSM-IV alcohol dependence is known to increase risk by at least twofold (2). Males are more likely than females to develop AUD (1–4), both within families of alcohol-dependent probands and in the general population (2).

In the United States, persons of African ancestry (AA) are less likely to develop AUD than persons of European ancestry (EA) (5–7). However, over recent years, there have been relatively greater increases of AUD in women and AA individuals compared with EA males (8).

Early age at first drink has been robustly associated with increased risk for adult drinking problems (9–11). There has been controversy over whether this is the best predictor of regular drinking and alcohol problems in adolescents/young

adults (11-14). However, it does appear to be associated with genetic vulnerability (15-18).

There is a known risk relationship between other psychiatric disorders and AUD. Adolescents with a mood disorder are at increased risk for alcohol problems and disorders (19–21) and vice versa (21). Scores on an internalizing scale were positively correlated with risk for alcohol and other drug use disorders (22,23). There is an extensive literature supporting the relationship of externalizing disorders (especially conduct disorder) to subsequent development of AUD (22,24–27). There is a substantial overlap between genetic vulnerability to AUD and vulnerability to externalizing disorders in general (28).

Genetic vulnerability to common disorders is quantifiable using polygenic risk scores (PRSs), which combine the cumulative effects of many genetic loci into a single metric

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CC BY-NC-ND license (http://creativecommons.org/licenses/by-nc-nd/4.0/). Biological Psychiatry: Global Open Science October 2022; 2:379–388 www.sobp.org/GOS (29–32). In practice, a PRS is calculated from a discovery genome-wide association study and applied to an independent target dataset. The PRS enumerates the number of copies of the risk allele carried by that individual at each locus (0, 1, or 2) and weights each risk allele by its effect size. Procedures for single nucleotide polymorphism (SNP) selection may generally be divided into 1) pruning and thresholding, which generates a set of scores based on different p value thresholds for the odds ratio estimates for disease association; or 2) Bayesian, which generates a single optimized score. Covariates for sex, ancestry, and other variables may also be added. Most of the literature on PRSs is based on thresholding methods, but Bayesian analysis is becoming more widely used because it obviates the need for multiple testing correction.

PRSs have been used in many complex medical conditions. Khera *et al.* (33) studied genetic risk for coronary artery disease, atrial fibrillation, type 2 diabetes mellitus, inflammatory bowel disease, and breast cancer. Using PRSs, they identified subgroups of subjects with 3 times, 4 times, and 5 times increased risk for each disorder. In a subsequent paper, Khera *et al.* (34) predicted severe obesity in a 20-year follow-up study of young adults, showing 1.3% severe obesity among subjects in the lowest PRS decile compared with 15.6% among subjects in the highest decile.

PRSs have been reported to discriminate between cases and controls with an area under the receiver operating characteristic (ROC) curve as high as 82% for schizophrenia, 65% for bipolar disorder, 58% for major depressive disorder, and 54% for anxiety, although there is substantial variability in estimates across studies (35). Recently Musliner *et al.* (36) showed a significant increase of conversion risk from major depressive disorder to bipolar disorder among subjects in different PRS subgroups in a sample of offspring of persons with bipolar disorder. Such studies have treatment implications (32).

In this paper, we explore the applicability of alcoholrelated PRSs to AUD in an adolescent and young adult sample. We have previously reported in the Collaborative Study of the Genetics of Alcoholism (COGA) dataset effects of sex, ancestry, family type (case/comparison), and comorbid psychiatric disorders on risk for early onset of AUD (37). This report extends this analysis to include PRS, the age of first drink, and peer drinking. All of these variables (with the exception of peer drinking and comorbidity) are known to be fixed in value prior to the onset of AUD in any subject. COGA subjects have participated in separate multicenter studies of PRSs (38,39), but not in combination with clinical variables.

METHODS AND MATERIALS

Subjects

This analysis was based on the COGA Prospective Study Dataset (22). Ascertainment sites were University of Connecticut (Farmington, CT), Indiana University School of Medicine (Indianapolis, IN), University of Iowa (Iowa City, IA), Washington University in St. Louis (St. Louis, MO), State University of New York Downstate (Brooklyn, NY), Howard University (Washington, DC), and University of California San Diego (San Diego, CA). All participants provided informed consent for study procedures. All protocols were approved by institutional review boards at the various institutions.

This sample (n = 2794) was a subset of the Prospective Study Dataset (N = 3286) (37); the subset included all dataset members with genotypic data. Adolescent and young adult subjects were assessed with a structured psychiatric interview [Semi-Structured Assessment for the Genetics of Alcoholism (40)] at 2-year intervals from 2004 to 2017. All subjects aged between 12 and 21 years at baseline assessment were invited to participate. Every subject with at least 1 complete interview (73% of those invited) was included in the dataset. The average age at first interview was 16.1 years (SD = 3.3), and the average age at last interview was 23.1 years (SD = 5.0). Most subjects (87%) came from extended families with a proband in treatment for an AUD (designated family type = case); 13% were from comparison families (family type = comparison), recruited from sources such as dental clinics and motor vehicle records. Subjects in comparison families were not prescreened and may be expected to have population rates of common disorders such as AUD (2). Subjects in the dataset were included in analyses regardless of initiation of alcohol use by the time of the baseline interview. More than 80% of subjects participated in at least one follow-up. The average number of interviews per subject was 4.0 (SD = 1.7). There was no evidence for selective attrition of affected subjects (22). All subjects had data on comorbid diagnoses, AUD diagnosis, peer drinking, and age of first drink. Genotyping was carried out and quality control procedures applied as previously described (41,42). Ancestry was assigned based on the first four genotypic principal components of population stratification (Figure S1). Individuals surrounding the CEU HapMap position were assigned to the EA sample, and those beyond a radius of about 0.002 units on the principal component 1 axis were assigned to the AA sample. To maintain power for familybased analyses, final ancestry assignment was based on the majority of individual family members. Families with an equal distribution were assigned to the most diverse group (AA). This analysis excluded subjects who were not assignable to either the EA or AA groups. There were 2794 subjects with genomewide association study data available for this analysis, 67% EA (1872 subjects from 559 families) and 33% AA (922 subjects from 219 families). Among the subjects, 51% were female and 49% were male.

The dependent variables for the analyses were age of onset (AOO) for all AUDs and severe AUD as defined by DSM-5 (43). Externalizing disorders and internalizing disorders were defined as in Nurnberger *et al.* (37). Briefly, we considered DSM-IV (44) diagnoses of externalizing disorders (attention-deficit/hyperactivity disorder, conduct disorder, antisocial personality disorder, oppositional defiant disorder, and drug use disorder excluding alcohol or tobacco) and internalizing disorders (major depressive disorder, panic disorder, obsessive-compulsive disorder, social phobia, and agoraphobia) if they occurred before or at the same time as the onset of AUD. Comorbidity was scored as membership in one of four groups: externalizing (at least one disorder), internalizing (at least one disorder), both, or neither.

Age of first drink was defined using the Semi-Structured Assessment for the Genetics of Alcoholism question "How old were you the first time you had your very first whole drink?" Peer drinking was defined using the Semi-Structured Assessment for the Genetics of Alcoholism question "When you were 12–17, how many of your best friends used alcohol?" Answers were none, a few, most, or all. For analysis, we divided responses into two categories: most friends drink and most friends do not drink.

Construction of PRSs

We used the Million Veteran Program (MVP) (45,46) datasets as discovery samples and examined EA and AA PRSs for ICD 9/10 AUD, Alcohol Use Disorders Identification Test-Concise (AUDIT-C) scores, and MAX_ALC, defined as the highest number of drinks a subject reported drinking during a single day in a typical month (46). The MVP dataset included AUD cases (56,000, including 34,000 EA and 17,000 AA) and controls (219,000, including 167,000 EA and 39,000 AA). This is the largest sample of AUD cases available currently (although there are larger samples with data on consumption or problematic use). Variants located within 500 kb of the index variant and having $R^2 > 0.25$ with the index variant were clumped. PRSs were calculated as the sum of allele counts weighted by the sign of the log odds ratio and the negative logtransformed p value for each SNP. The weighting by p value was used because it is robust to variations in sample size from SNP to SNP. Each PRS was tested at nine thresholds. A p value of 3.3 imes 10⁻⁴ was considered significant for a PRS after Bonferroni correction (although this is probably conservative, because PRSs are correlated). Calculations were performed with PLINK (47). For comparison, we also ran PRS-CS (48) for EA subjects and PRS-CSx (49) for AA subjects. For these analyses, the Bonferroni correction was 1.6 \times 10 $^{-2}.$ All PRS results were standardized (mean = 0, SD = 1).

We applied PRS to the COGA Prospective Study EA dataset, controlling for sex, ancestry (using the first four principal components as indicated by the scree plot) (Figure S1), family history, and relatedness. We used the frailty model to capture relatedness for the COGA data; this is the most widely used method to handle correlated survival data (50). Specifically, the frailty model puts a random effect term in the Cox regression formula such that the members of the same family will share the same random effect.

PRS analyses were performed separately in AA subjects using PRSs derived from the MVP AA dataset. The relatively small number of subjects (4%) with non-EA, non-AA ancestries were not included in PRS analyses.

Kaplan-Meier curves were used to estimate the survival function of AOO of AUD. A Cox proportional hazards model was used to test the relationship of AOO with PRSs adjusted for sex, relatedness, and case/comparison family status in the original model. Other variables added into the model for specific analyses included comorbid disorders, age of first drink, and peer drinking. The proportional hazards assumption was tested in these analyses and was not violated. Time-dependent ROC curves were created to examine predictive value for the diagnosis of any AUD or severe AUD at a given age (and an integrated value for the age range 15–27). Area under the curve (AUC) was calculated to assess the performance of the model. All statistical analyses (except calculation of PRS) were run using SAS/STAT 15.1 (51).

RESULTS

By the end of the follow-up period, 1544 of 2794 subjects were unaffected, and 606 were diagnosed with mild AUD, 365 with moderate AUD, and 279 with severe AUD. In total, 45% of subjects were diagnosed with AUD during the assessment period [for additional detail on outcomes, see (38)]. Table 1 shows the distribution of subjects in various clinical categories.

Results of the Cox proportional hazards model are presented in Tables 2 and 3. Risk was increased for males, subjects from case families, and subjects with increased PRS. PRS from MAX_ALC.P1 (which uses the threshold p < .1) was associated ($p < 6.1 \times 10^{-6}$ by Cox test) with AOO of AUD in EA subjects (Table 1). PRSs derived from AUD or AUDIT-C scores did not show a strong relationship with AOO in EA subjects (p > .001 at most thresholds). PRS variables were not significantly associated with AOO in AA subjects (Tables S1A and S1B).

A histogram of PRS in EA subjects from case and comparison families is shown in Figure 1. The distribution of PRS is shifted to the right in case subjects, although there is substantial overlap. The shift is not significant in the thresholded analysis ($p = 4.43 \times 10^{-4}$), but it is when calculated by PRS-CS ($p = 2.89 \times 10^{-6}$). A similar histogram for AA subjects is shown in Figure S2 (no significant shift by thresholded analysis or by PRS-CSx).

In general, recalculation of PRS by PRS-CS and PRS-CSx produced more significant p values (or closer to significance in the case of AA subjects) but did not change the results substantially, except as noted above.

Figure 2 illustrates the relationship between PRS and onset of AUD in EA subjects ($p = 5 \times 10^{-6}$ by the Cox model; $p = 5 \times 10^{-7}$ by log-rank test). The median AOO for AUD for subjects with a PRS in the top half of the distribution is 20, while the median AOO for subjects in the lower half of the PRS distribution is 22.

Combining PRSs With Demographic and Clinical Variables

As noted above, PRS was significantly associated with AOO covarying for sex, family type, and ancestry (Table 2). Adding in comorbid disorders, age of first drink, and peer drinking, PRS MAX_ALC.P1 remained associated with AOO of AUD in EA subjects (hazard ratio = 1.14, Cl = 1.06-1.22, $p = 2.11 \times 10^{-4}$) (Table 3). For data in AA subjects, see Tables S1A and S1B. There is no significant relationship between PRS and AOO in AA subjects.

Risk Calculation Using ROC Curves

AUC peaks at 88% (all AUD) and 95% (severe AUD) for EA subjects (Table S2). Using a cut-point of 0.5, sensitivity is

	European Ancestry, $n = 559$ Families			African Ancestry, $n = 218$ Families			Total
Variable	n	Column% ^ª	Row% ^b	п	Column% ^ª	Row% ^b	n (%)
Sex							
Female	953	50.91%	66.74%	475	51.52%	33.26%	1428 (51.11%)
Male	919	49.09%	67.28%	447	48.48%	32.72%	1366 (48.89%)
Total	1872	100.00%	67.00%	922	100.00%	33.00%	2794 (100.00%)
Comorbidity							
Internalizing only	89	4.75%	80.18%	22	2.39%	19.82%	111 (3.97%)
Externalizing only	508	27.14%	59.42%	347	37.64%	40.58%	855 (30.60%)
Both	170	9.08%	66.93%	84	9.11%	33.07%	254 (9.09%)
Neither	1105	59.03%	70.20%	469	50.87%	29.80%	1574 (56.34%)
Total	1872	100.00%	67.00%	922	100.00%	33.00%	2794 (100.00%)
Peer Drinking							
Peer drinking	858	45.83%	70.04%	367	39.80%	29.96%	1225 (43.84%)
No peer drinking	1012	54.06%	64.58%	555	60.20%	35.42%	1567 (56.08%)
Not available	2	0.11%	100.00%	0	0.00%	0.00%	2 (0.07%)
Total	1872	100.00%	67.00%	922	100.00%	33.00%	2794 (100.00%)
Case/Comparison Status							
Subjects from case families	1559	83.28%	64.24%	868	94.14%	35.76%	2427 (86.86%)
Subjects from comparison families	313	16.72%	85.29%	54	5.86%	14.71%	367 (13.14%)
Total	1872	100.00%	67.00%	922	100.00%	33.00%	2794 (100.00%)
Age of First Drink							
Has onset age of first drink	1674	89.42%	67.61%	802	86.98%	32.39%	2476 (88.62%)
No first drink reported	198	10.58%	62.26%	120	13.02%	37.74%	318 (11.38%)
Total	1872	100.00%	67.00%	922	100.00%	33.00%	2794 (100.00%)
AUD Diagnosis							
Mild	423	22.60%	69.80%	183	19.85%	30.20%	606 (21.69%)
Moderate	279	14.90%	76.44%	86	9.33%	23.56%	365 (13.06%)
Severe	223	11.91%	79.93%	56	6.07%	20.07%	279 (9.99%)
No AUD	947	50.59%	61.33%	597	64.75%	38.67%	1544 (55.26%)
Total	1872	100.00%	67.00%	922	100.00%	33.00%	2794 (100.00%)

Table 1. Description of the Sample by Clinical/Demographic Variables

AUD, alcohol use disorder.

^aColumn% = % of total in that column.

 ${}^{b}Row\% = \%$ of total in that row.

72% and specificity is 87% at age 15. Sensitivity is 73% and specificity is 67% at age 20. In most models, peak AUC is achieved with inclusion of all variables. AUC power is primarily driven by the variable age of first drink, and the independent effects of other variables (including PRS) are generally small. PRS adds 8.6% to AUC by itself and 0.7% after all clinical variables in the integrated EA all AUD

analysis; in the integrated EA severe AUD analysis, PRS adds 10.4% by itself but adds no additional variance following all clinical variables. AUC peaks at 88% (all AUD) and 92% (severe AUD) for AA subjects (Table S2). In the AA analyses, PRS adds 5.5% initially to all AUD and 13.5% initially to severe AUD but <1% when added after all clinical variables.

Table 2. Cox Proportional Hazards Model for Sex, PRSs (EA_MAX_ALC.P1), and Family Type (Case/Comparison) in EA Subjects: Analysis of Maximum Likelihood Estimates

Parameter	Parameter Estimate	Standard Error	χ ² 1	p Value	HR	95% Confidence Limits
Sex	-0.28030	0.06852	16.7344	$4.30 imes10^{-5}$	0.756	0.661-0.864
EA_MAX_ALC.P1 (With PC1-PC4)	0.16451	0.03600	20.8803	$4.89 imes10^{-6}$	1.179	1.099–1.265
Family Type	0.54092	0.11717	21.3110	$3.90 imes10^{-6}$	1.718	1.365–2.161

The model is also adjusted for ancestry using the first four PCs of population stratification (PC1–PC4). All HRs show the effect of a particular variable after accounting for the effects of all other variables.

EA, European ancestry; HR, hazard ratio; PC, principal component; PRS, polygenic risk score.

Parameter	Parameter Estimate	Standard Error	χ ² 1	p Value	Hazard Ratio	95% Confidence Limits
Sex	-0.20281	0.06854	8.7550	$3.09 imes10^{-3}$	0.816	0.714-0.934
EA_MAX_ALC.P1 (With PC1-PC4)	0.13171	0.03554	13.7311	$2.11 imes 10^{-4}$	1.141	1.064–1.223
Family Type	0.16013	0.11287	2.0128	$1.56 imes10^{-1}$	1.174	0.941-1.464
Age of First Drink	-0.15056	0.01239	147.6970	$5.53 imes10^{-34}$	0.860	0.840-0.881
Peer Drinking	0.68893	0.07471	85.0399	$2.92 imes10^{-20}$	1.992	1.720-2.306
Comorbidity						
Both	0.30369	0.10633	8.1578	$4.29 imes10^{-3}$	1.355	1.100-1.669
Externalizing only	0.32966	0.07749	18.0974	2.10×10^{-5}	1.390	1.195–1.619

Table 3. Cox Proportional Hazards Model for Sex, PRSs (EA_MAX_ALC.P1), Family Type (Case/Comparison), Age of First Drink, Peer Drinking, and Comorbidity in EA Subjects: Analysis of Maximum Likelihood Estimates

The model is also adjusted for ancestry using the first four PCs of population stratification (PC1–PC4). All hazard ratios show the effect of a particular variable after accounting for the effects of all other variables.

0.5442

 4.61×10^{-1}

0.15937

EA, European ancestry; PC, principal component; PRS, polygenic risk score.

-0.11756

Subgroups Based on Divisions of the PRS Distribution

Internalizing only

Because AUC calculations do not capture the power of PRS at the extremes of the distribution, we plotted AOO for AUD in quartiles of EA PRS (Figure 3). Median AOO for subjects in the four quartiles were 20, 20, 22, and 24, respectively. By age 25, 72% of subjects in the highest quartile manifested AUD compared with 54% in the lowest quartile.

0.889

0.651-1.215

Figure 4 shows an analysis of AOO for severe AUD in EA subjects using PRS quartiles. By the age of 25, 35% of subjects in the top quartile were diagnosed with severe AUD as compared with 16% of subjects in the bottom quartile. Subgroup analyses in AA subjects are presented in Figures S2–S5.

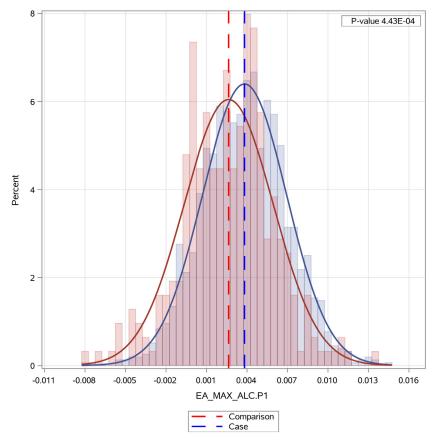


Figure 1. Polygenic risk score distribution in subjects from European ancestry (EA) case families (blue) and comparison families (red). The polygenic risk score distribution is shifted to the right in the case families, adjusting for sex and principal component 1 through principal component 4. The shift is not significant in the thresholded analysis (shown here for consistency) but is significant when calculated by the PRS-CS method (48) ($p = 2.89 \times 10^{-6}$). EA_MAX_ALC.P1, PRS based on MAX_ALC in European ancestry subjects using the threshold p < .1.

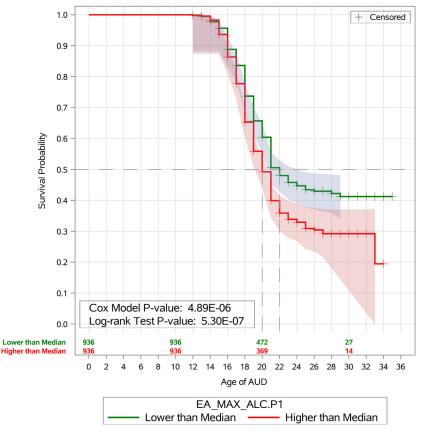


Figure 2. Age of onset for first diagnosis of alcohol use disorder (AUD) in European ancestry subjects with high polygenic risk scores (using a median split) is lower than age of onset for AUD for subjects with low polygenic risk scores. The Cox model shows the *p* value incorporating sex and principal components for ancestry ($p = 4.89 \times 10^{-6}$). The log-rank test is not adjusted for covariates ($p = 5.30 \times 10^{-7}$). EA_MAX_ALC.P1, PRS based on MAX_ALC in European ancestry subjects using the threshold p < .1.

DISCUSSION

Association of PRSs With Onset of AUD

In our analysis, a PRS derived from MAX_ALC was associated with AOO of AUD in our high-risk sample. Clinical prediction that includes easily measurable variables, such as comorbid conditions, age of first drink, and peer drinking, shows efficacy (0.88–0.95) that approaches the range of clinical utility (52), with or without PRS. The predictive value of PRS is most evident when the extremes of the PRS distribution are compared, as illustrated by a 1.3-times increase in the probability of developing any AUD and a 2.2times increase in the probability of developing severe AUD by age 25 among EA individuals from the top and bottom PRS quartiles.

PRSs may be presumed to increase in predictive accuracy with the size of the discovery sample. We used the largest sample available at this time (MVP). Because AUD is a common disorder in the population [>29% lifetime prevalence (8)] and is only moderately heritable [approximately 55% (53)], we may expect that very large samples will be necessary to achieve optimal prediction (54).

PRSs and Ancestry

In our Cox model analyses, PRS was not significantly associated with AOO of AUD in the AA sample (although predictive ability using ROC was similar in the two groups using primarily clinical variables). Recent studies (55) show that PRSs for bipolar disorder derived from EA subjects are reasonably predictive when applied to East Asian subjects but less so when applied to AA populations. It is well established that linkage disequilibrium blocks (units of correlated genetic markers) are smaller in AA subjects than in other populations. This is related to the history of the human species, which extends an order of magnitude longer on the African continent than in other areas of the world (56). Thus, there are different allele frequencies and greater variation in allele frequency in AA subjects compared with those of other ancestries. The reliability of PRS developed from other ancestral groups is consequently less in AA subjects. Because PRS would currently be less useful for AA subjects than for EA subjects, questions arise regarding inequities in clinical application (57). One solution is increased emphasis on collection of genetic samples from diverse populations, especially those of AA. We should note that the sample sizes in both the discovery (MVP) and target (COGA) datasets in these analyses are considerably smaller for AA subjects compared with EA subjects.

Clinical Utility of Predictive Algorithms

Effective treatments for AUD are now available, such as cognitive behavioral therapy, oral naltrexone, and long-acting

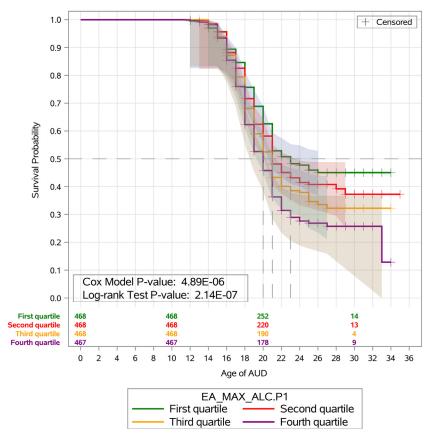


Figure 3. Kaplan-Meier curves (survival analysis) for age of first alcohol use disorder (AUD) diagnosis for European ancestry subjects in quartiles of the polygenic risk score distribution. Shaded areas around each curve represent 95% confidence intervals. The Cox model incorporates sex and ancestry principal components. The log-rank test does not incorporate covariates. EA_MAX_ALC.P1, PRS based on MAX_ALC in European ancestry subjects using the threshold p < .1.

injectable naltrexone, among others (58,59). These treatments are capable of saving lives, relationships, and careers, as well as minimizing medical comorbidities. Specific treatment and/or monitoring might be considered for adolescents and young adults in high-risk subgroups who are already drinking and manifesting problematic alcohol use patterns, such as binge drinking.

We note the importance of age of first drink in the AUC analyses. Early exposure to alcohol may have causal effects on later addictive behaviors. Adolescent exposure to alcohol preferentially increases alcohol drinking during adulthood in rat models (60), along with brain changes in important neuro-transmitter systems and brain areas for appetitive behavior (61,62). Similar mechanisms in humans are a plausible hypothesis (63,64). Prevention strategies to delay exposure to alcohol might receive additional attention in public health efforts at harm reduction.

Limitations

The discovery sample differed from the target sample in several important respects. MVP is a study of older adult veterans, predominantly male. The COGA sample is mixed male and female adolescents and young adults and is a sample of subjects at risk for AUD. Perhaps one consequence of this is the observation that PRS for MVP AUD was less effective in the COGA sample than PRS for MAX_ALC; types of AUD diagnoses in the two samples may differ more than this quantitative lifetime measure. Replication of these results in independent cohorts would help establish their generalizability.

When discussing predictive testing, it is important to consider the issue of stigmatization, especially in young people. We are not developing algorithms for clinical use in individuals below drinking age at this time. However, we would advocate consideration of algorithms for prediction of early-onset AUD when problematic drinking is identified in adolescents. The fact that the algorithm includes genetic information is no reason to consider it additionally problematic in terms of stigmatization because heritable disorders may be successfully treated. Concerns regarding genetic information should be addressed as aspects of public health education, directed at consumers, providers, and ultimately the general population (65). The larger issue at present is difficulties in interpretation of individual PRS values, especially in mixed ancestry populations.

Methodologic improvements of PRS [e.g., PRS-CS, included here for comparison (48)], which obviates correction for multiple testing, and the use of SNP weights based on gene expression in PRS (66) might provide improved predictive capacity.

It is important to realize that physicians and scientists are no longer the only gatekeepers for individual genetic information.

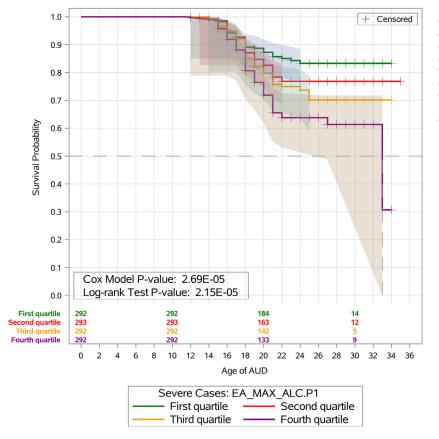


Figure 4. Kaplan-Meier curves (survival analysis) for age of first severe alcohol use disorder (AUD) diagnosis for European ancestry subjects in quartiles of the polygenic risk score distribution. Shaded areas around each curve represent 95% confidence intervals. The Cox model incorporates sex and ancestry principal components. The log-rank test does not incorporate covariates. EA_MAX_ALC.P1, PRS based on MAX_ALC in European ancestry subjects using the threshold p < .1.

Direct-to-consumer companies offer genome-wide data and, in some cases, PRSs for multiple conditions, including psychiatric disorders. Individual patients may obtain this information and may bring it to their doctors. It is incumbent on clinical professionals to guide patients on conservative interpretation of such data.

Clinical trials may eventually be used to formally test the value of PRSs in combination with appropriate clinical variables. The model might be similar to that used in recent trials using pharmacogenetic testing (67), although appropriate follow-up would presumably be more extended. This would provide an opportunity for real-life examination of the feasibility and utility of such scores. Appropriate clinician training would be an essential part of such trials.

In conclusion, several variables had significant effects on AOO of AUD in this study: sex, family history, age of first drink, peer drinking, comorbidity, and PRS. Discriminatory power in the ROC model was maximized by using age of first drink along with other variables. PRS was useful in identifying subgroups at unusually high risk. Such algorithms might have a place in the future clinical practice of psychiatry. Larger samples, especially AA samples, will be necessary to support more effective use of PRS in diverse clinical populations. In combination with clinical variables, PRS may aid in prediction of outcome and clinical decision making. Following additional study, clinical trials may help to assess feasibility and utility.

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