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Collaborative HRM, climate for cooperation, and employee intra-organizational social ties in high-technology firms in China: A cross-level analysis

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Individual social ties have been an important source of competitive advantages for hightech firms in the knowledge economy. However, the existing cross-level studies have mainly investigated the impact of HRM systems on traditional individual attitudinal or behavioral outcomes, and few studies have examined the effect of SHRM on individual social ties. Based on the data collected from 363 knowledge employees working in 64 high-tech firms in China, we examine the cross-level relationships among collaborative HRM practices, climate for cooperation and employee intra-organizational social ties. The hierarchical linear model results show that organizational-level collaborative HRM practices have significant positive effects on the number and strength of individual-level intra-organizational social ties, and the climate for cooperation mediates the positive cross-level relationship between collaborative HRM and individual intra-organizational social ties. This study makes three contributions to the literature. First, with a cross-level model, this study helps us better understand how collaborative HRM acts as an approach to manage individuals' social capital formation. Second, this study makes contribution to the social network literature by showing how organizational contextual factors (HRM practices and organizational climate) affect employee individual social ties. Third, based on the AMO model, this paper developed a more clear construct and a three-dimension measurement of the collaborative HRM.

KEYWORDS

collaborative HRM, climate for cooperation, employee social ties, cross-level analysis, hightechnology firms

Introduction

Over the past two decades, scholars have widely recognized the essential role that strategic human resource management (SHRM) plays in driving organizational effectiveness (Huselid, 1995; Combs et al., 2006). Recently, the cross-level effects of the organizational HRM systems (e.g., a bundle of favorable HR practices) on individual employee outcomes have received an increasing amount of attention (Huang et al., 2018; Pahos and Galanaki, 2022). However, the existing cross-level studies have mainly investigated the impact of HRM systems on traditional individual attitude or behavior varialbes, such as job satisfaction, organizational commitment, and organizational citizenship behavior (Kehoe and Wright, 2013; Heffernan and Dundon, 2016; Zhang et al., 2018), and few studies have examined the effect of SHRM on individual social ties.

Individual social tie (or relational tie) is the social links between actors (Hollenbeck and Jamieson, 2015), which is the core of a social network (Seibert et al., 2001). With the advent of the knowledge

economy, employee social ties have been an important source of competitive advantages for knowledge intensive companies since the processes of acquiring, transforming and integrating valuable knowledge often occurs in individual social interactions (Nahapiet and Ghoshal, 1998; Kale et al., 2000; Kang et al., 2007). Empirical studies have revealed that social interactions in social networks can influence knowledge exchange and knowledge creation, and ultimately lead to better team and organizational effectiveness (Hansen et al., 2005; Chen and Huang, 2007).

Scholars have highlighted the roles of HRM practices in influencing the way employees interact with their colleagues (Brass et al., 2004; Kaše et al., 2009). For instance, HR practices communicate organization's rules, procedures, and policies, which are critical in forming employee social ties (e.g., Liebeskind et al., 1996; Leana and Van Buren, 1999). Moreover, HR practices such as extensive training and work design for example provides employees with the structural connections to develop constructive interpersonal interactions, which ultimately lead to firm competitive advantages (Evans and Davis, 2005; Lengnick-Hall and Lengnick-Hall, 2006). Accordingly, several empirical studies have investigated the relationships between the collaboration-based or relation-oriented HRM systems and collective outcomes. For example, Chuang et al. (2016) examined an "HRM system for knowledge-intensive teamwork" and demonstrated its impact on the team-level knowledge generation processes of R&D teams. Kehoe and Collins (2017) reported that a set of relationshiporiented HR practices improved the performance of units by increasing unit members' aggregate access to knowledge within and outside units. Nieves et al. (2016) reported that collaborative HRM are an antecedent of product innovation. Colakoglu et al. (2021) argued that the perceived collaboration-based HR systems promote information exchange. Chen Y. Y. et al. (2011) found a positive interaction effect of collaborative HR practices reported by employees and network range on objective sales performance. Zhou et al. (2012) proposed that the application of collaboration-based HRM will be positively related to external talent deployment to achieve innovation and accentuate firm performance. However, so far, there is a lack of studies to examine whether and how organizational-level HRM practices will influence individual-level employee social ties.

In this study, we propose that the collaborative HRM, as an HRM configuration investing in interpersonal relationships rather than in individuals (Lepak and Snell, 1999), focusing on permeable and network intimate work structures, team development, and group incentives and emphasizes co-operation, information sharing and knowledge transfer (Lepak and Snell, 1999; Youndt and Snell, 2004), is an important approach to promoting individual social ties. According to the AMO model (Appelbaum et al., 2000; Jiang et al., 2012), we argue that collaborative HRM can directly affect employee social network by enhancing individuals' ability, motivation and opportunity to build social ties with their colleagues. Furthermore, based on the implication of previous studies that HRM systems usually influence employee outcomes through building a social context (Ferris et al., 1998; Collins and Smith, 2006), we propose that an organizational climate for cooperation will mediate the relationships between collaborative HRM and the number and strength of employees' intra-organizational social ties.

The number and strength of employees' social ties are two frequently-used indicators to describe a social network (Senaratne et al., 2021). The number of social ties is usually defined as the number of relationships maintained by individuals in a social network, and it indicates individuals' influence and power within the social network (Müller-Prothmann, 2007; Borgatti et al., 2013). The strength of social ties describes the degree of connectivity such as the likelihood of information flows among individuals. When the strength is high, individuals are motivated to provide information and support to others (Kim et al., 2011; Borgatti et al., 2013).

Our research model is shown in Figure 1. As illustrated, we applied the cross-level mediation-upper mediator model, also known as the 2–2-1 model (Zhang et al., 2009). In this model, collaborative HRM and organizational climate for cooperation are variables at the organizational level, and the number and strength of employees' intra-organizational social ties are variables at the individual level. We collected data from both organizations and employees, and conducted the cross-level analysis with Mplus. We predict that the level-2 independent variable (i.e., collaborative HRM) influences the level-1 dependent variables (i.e., tie number and tie strength) through the level-2 mediating variable (i.e., climate for cooperation).

Our research will make three contributions. First, with a cross-level model, this study helps us better understand how collaborative HRM acts as an approach to manage individuals' social capital formation. This not only expands the literature on strategically-targeted HRM with a multilevel perspective (Lepak et al., 2006; Takeuchi et al., 2009), but also unfolds the mechanisms that link HRM and employee outcomes. Second, this study makes contribution to the social network literature by showing how organizational contextual factors (HRM practices and organizational climate) affect employee individual social ties. Past research has focused on the individual level antecedents of social ties, such as individuals' gender, age, and self-monitoring levels (Brass et al., 2004; Osman-Gani and Rockstuhl, 2009), however few studies have investigated organizational antecedents (Kaše et al., 2009). From the SHRM perspective, we demonstrate that organizational HRM practices can be effective tools to develop employee social ties. Third, based on the AMO model, this paper developed a more clear construct and a three-dimension measurement of the collaborative HRM, since previous studies were not so consistent and detailed in the structure and measurement of the collaborative HRM (Lepak and snell, 2002; Lopez-Cabrales et al., 2009).

Theoretical background and hypotheses

Collaborative HRM

The pioneer literature studies on collaborative HRM are primarily concerned with information sharing and building cooperative relationship between collaborators, such as enterprises and their suppliers (Lepak and Snell, 1999, 2002). As Lepak and Snell (1999, p. 41) stated, "although alliances may involve structural arrangements in which employees from both parties work together, HR systems that encourage and reward cooperation, collaboration, and information sharing are also likely to be necessary." Accordingly, they argued that organizations use team-building training, communication mechanisms, and the like to facilitate information sharing and the transfer of knowledge. Later on, the concept of collaborative HRM gained relevance and applicability within organizations. Lopez-Cabrales et al. (2009) argued that collaborative HRM (including teamwork design, teamwork skills-based training, socialization programs, team-based appraisal systems, and collective reward systems) are critical for disseminating specialized knowledge within organizations. Specifically, as a kind of HRM with specific goal orientation, collaborative HRM comprises of a set of HRM practices that promotes interaction and collaboration among employees.



For example, Kang et al. (2007) and Lopez-Cabrales et al. (2009) argued that collaborative HRM practices (including teamwork design, job rotation, teamwork skills based selection and training, socialization programs, team-based appraisal systems and collective reward systems) are critical for disseminating specialized knowledge within organizations.

We adopted the widely used ability-motivation-opportunity approach (e.g., see also Combs et al., 2006; Lepak et al., 2006; Jiang et al., 2012) and argue that the collaborative HRM should include ability-enhancing practices, such as employee selection and promotion that prioritize candidates with strong collaborative competency, and workplace mentoring system that binds employees into learning partnership; motivation-enhancing practices such as performance evaluations that gives a substantial weight to employees' social skills, merit pay increase that is driven by employees' teamwork skills, and a reward system that actively recognize employees' performance in collaborative tasks; opportunityenhancing practices such as rotating jobs, sponsoring employees' social gatherings, and wide utilization of interdependent work design. Employees who have the ability to cooperate with each other do not necessarily have the motivation to do so, and employees who have both the ability and motivation do not necessarily have the opportunity to cooperate with their colleagues. Consequently, only when all three dimensions are available can social ties be more generated at the individual level.

Employee intra-organizational social ties

A social network is defined as a relatively stable system of social ties linking a defined set of persons or social actors (Burt, 1997; Tsai and Ghoshal, 1998). Individual social ties are the core of a social network (Seibert et al., 2001). In the organization context, employees' intraorganizational social ties refer to the social relations that employees have with their colleagues within the organizations, which can be characterized by tie number and strength. More specifically, the tie number conveys how many peers the employee interacts with for information and knowledge exchange, while the tie strength conveys how often and tightly the employee interacts with their colleagues.

Studies have revealed that a greater number of ties help employees obtain a broader range of perspectives (Dean and Brass, 1985), enabling

employees to make more extensive use of new information (e.g., policy, technology) introduced in their workplace. In addition, frequent interaction with colleagues helps employees gather information quickly, and close ties with others allow individuals to dispense with formality and self-censorship, and to get to the heart of issues (Pil and Leana, 2009). Besides, since every member of an organization is embedded in a wide and complex social network, when individuals maintain a strong social network (e.g., a social network with high a tie number and tie strength), they are more well connected to the larger social network within the organization, and can be more effective to obtain diversified information (Sosa, 2011; Hirst et al., 2015). Prior studies have shown that strong social interactions or social networks facilitate knowledge sharing and creation (Chen and Huang, 2007), which lead to high effectiveness for both individuals and organizations (Coleman, 1988; Leana and Pil, 2006; Oh et al., 2006).

The importance of individual social ties on employee performance such as innovation performance and creativity has been underscored in hightech enterprises (Hu, 2008; Zhou et al., 2009). In these firms, many professionals are heavy users of tacit knowledge, meaning that the knowledge needed by these individuals are not obtained readily from an organization's formal documents (Liu and Liu, 2008). Relatedly, when they encounter difficult technical or professional problems, they usually first seek assistance and support from their colleagues. Moreover, these employees usually exchange the knowledge and resources to those with high-quality interpersonal relationships colleagues (Bouty, 2000). Consequently, highquality social ties in high-tech firms is an important factor that employees to integrate resources, exchange information and gain support, which all contribute to their individual performance. Nevertheless, few studies have explored the antecedents of social ties within an organizational context.

Collaborative HRM and employee intra-organizational social ties

We propose that organizations can implement collaborative HRM practices to enlarge and strengthen employees' social ties with their colleagues. HRM practices such as selecting job candidates based on their collaborative ability, providing training programs focusing on interpersonal relation building skills, and implementing mentoring systems can improve employees' abilities to build social ties with their colleagues. HRM practices like teamwork-based performance appraisal and compensation design, promotion based on the ability to work with others can motivate employees to have interaction with more colleagues in a more often and closer manner. Organizations can also provide opportunities for employees to interact with their colleagues by adopting HRM practices such as providing time and resources for internal social activities, designing an interdependent working structure, implementing job rotation. Taken together, when organizations implement collaborative HRM practices, they foster the abilities, motivation, and opportunities necessary for social tie formation on the employees side.

We further argue that collaborative HRM practices influence employees' intra-organizational social ties through three mechanisms, namely, the assortative, relational and proximity mechanisms (Rivera et al., 2010). First, collaborative HRM help employee develop social ties through an assortative mechanism. HRM practice such as selecting job candidates based on their collaborative abilities, providing training focusing on teamwork skills will increase employees' common trait of collaborative spirit and abilities. Research has shown that individuals tend to foster close relationships with similar others (Brass et al., 2004). Consequently, through the assortation due to the implementation of HR practices, employees can increase the chances to interact with similar others in the form of high tie numbers, and also increase the likelihood to develop strong social interactions in the form of high tie strength.

Second, collaborative HR practices activate a relational mechanism of network building. Collaborative HR practices facilitates employees' social tie formation in the form of the tie number. HR practices such as team-based design, job rotations, provide formal opportunities for employees to interact with other colleagues. Informally, training, and company social events also provide more opportunities for employees to socialize. Moreover, the appraisal, rewarding, and promotion practices emphasizing collaborative results further encourage employees to work together. Moreover, collaborative HR practices further promote employees develop high tie strength. This is because, with the opportunities to work together, employees are likely to foster high-quality social tie as a consequence of reciprocation and acquaintance. Past research has documented that individuals tend to develop meaningful social ties in a long run when favors are reciprocated during interactions, and when work with those they have worked with in the past (Guimera et al., 2005).

Third, collaborative HRM practices facilitate employees to develop their social tie through a proximity mechanism, which means that individuals tend to formulate social ties when they are located within a short distance either in terms of physical space or social foci (Rivera et al., 2010). Interdependent work design and information sharing meetings give employees more opportunities to be closely located in the workplace. Sponsoring company social events and informal organizations allows employees to meet together more often in the same physical space and concurrently reduces the social distances among them. Team-oriented appraisal and reward systems also direct employees' interests toward their common goals and make them feel close with each other socially. Similar to the abovementioned relational mechanism, these opportunities for coordinate and work together can lead to strong social interactions in terms of both tie number and strength. Based on the above theoretical arguments, we propose the following hypothesis:

Hypothesis 1a: Collaborative HRM is positively related to the number of employees' intra-organizational social ties.

Hypothesis 1b: Collaborative HRM is positively related to the strength of employees' intra-organizational social ties.

Climate for cooperation as the mediator

Organizational HR systems can build a strong social context to guide employees' expectations and behaviors (Ferris et al., 1998; Bowen and Ostroff, 2004; Collins and Smith, 2006). Researchers use the concept of organizational climate to describe such social contexts. Organizational climate can be defined as organizational members' perception of formal and informal organizational policies, practices, and procedures (Schneider et al., 1998). Lepak et al. (2006, p. 224) argued that, "Organizational climate has been positioned as a key intermediate variable between organizational context and work outcomes. Specifically, organizational practices, policies, and procedures are argued to influence organizational climate, while organizational climate influences employees' collective attitudes and behaviors, which in turn influence organizational effectiveness." As a result, we propose that the organizational climate may play a mediating role in the relationship between the collaborative HRM and employees' behaviors relating to social interactions. Following past research in SHRM, instead of using a global climate concept, we focus on the climate for cooperation to represent the specific focus of strategic focus of the organization (Schneider et al., 1998; Lepak et al., 2006). Collins and Smith (2006) defined a climate for cooperation as the organizational norms that emphasize personal efforts toward group outcomes rather than individual outcomes. Nahapiet and Ghoshal (1998) argued that a climate for cooperation is a norm that significantly influences the exchange process, which opens the channel for other groups to exchange knowledge and determine the motivation for such an exchange. Thus, we consider a climate for cooperation as the collective norm perceived by employees, in which employees are expected and rewarded to communicate and cooperate with others.

Following Lepak et al. (2006), we propose that the positive relationships between the collaborative HRM and employee intraorganizational social ties (tie number, and tie strength) are mediated by the climate for cooperation. Ample research has shown that HRM can create an organizational climate that elicits certain behaviors from employees (Ferris et al., 1998; Collins and Smith, 2006; Lepak et al., 2006). Similarly, collaborative HR practices, such as promotion based on the ability to work with others, training on intra-firm relationship-building and selecting job candidates on the basis of their ability to collaborate in teams, interact with each other, consistently sending clear signals to employees that internal exchange and interactions are expected and rewarded by the organizations, thus contributing to the shared beliefs and perceptions regarding cooperative norms. In other words, when organizations adopt collaborative HR practices, the employees will observe a strong climate for cooperation.

When the norm of cooperation is established, it is natural to expect employees to interact with their colleagues more extensively and more frequently, since cooperation and internal exchange will be expected and rewarded by the organizations. According to the social identity theory, employees' social and work behaviors should be socialized under the influence of the common values of the organization (Louis, 1980). Expectedly, members within an organization would behave according to the common characteristics of the organization and make their cognition and behavior consistent with other members in the organization. Therefore, the cooperation climate would guide employees to actively communicate and interact with other members of the organization in order to maintain the identity of members of the organization. This will increase employees social tie numbers and strength. On the contrary, if the cooperative climate is weak or nonexistent, employees would observe that the organization has lower requirements on interaction with colleagues, and thus employees' behavior of building social connections would be drastically reduced (Chen and Huang, 2007).

Moreover, a climate for cooperation also promotes the formation of employee social networks by reducing the sense of competition among the employees. For knowledge workers, knowledge is a source of power and job security (Davenport and Prusak, 1998), therefore, they tend to have reserved attitudes towards sharing knowledge and information with their colleagues, which undermine the sharing and diffusion of knowledge (Reagans and McEvily, 2003; Collins and Smith, 2006). However, a climate for cooperation can reduce the sense of competition among employees by motivating them to focus more on the cooperation than on individual performance (Ingram and Roberts, 2000). Consequently, employees will be encouraged to interact with their colleagues and share what they know. Therefore, we propose the following hypothesis:

Hypothesis 2a: Climate for cooperation mediates the positive crosslevel relationship between collaborative HRM and the number of employees' intra-organizational social ties.

Hypothesis 2b: Climate for cooperation mediates the positive crosslevel relationship between collaborative HRM and the strength of employees' intra-organizational social ties.

Materials and methods

Sample and research procedures

The data for this study were elicited from online questionnaire surveys which were distributed to information technology (IT) companies in *Zhongguancun*, Beijing, China. This survey was conducted with the assistance of the *Beijing Zhongguancun Association of IT Professionals*. Due to the multiplicity of employee groups and subsequent multiplicity of HR systems within a firm (Lepak and Snell, 1999, 2002), our survey only focused on core knowledge employees whose primary duties and responsibilities are mainly related to R&D departments, engineering project<u>s</u>, and related departments. We ensured that the data for collaborative HRM were answered by both line managers and human resource managers, while the data of climate for cooperation and employee social ties were answered by core knowledge employees.

Overall, 126 out of the 213 firms that agreed to participate in the survey returned the questionnaires (59.1% return rate). Consequently,

126 HR managers, 287 line managers, and 548 employees completed the questionnaire. We included only firms that provided data for HR managers, line managers, and at least three core knowledge employees, thus reduced the number of qualified firms to 64, with 5.7 employees (range 3–17), 3.3 line managers (range 1–13) and 1 HR manager per firm on average. The sampled firms were, on average, 11.8 year old (sd = 7.1, min = 2, max = 38), hiring 815 employees (sd = 1,115, min = 44, max = 5,272). Furthermore, among the sampled core knowledge employees, 56.0% were male and 44.0% were female, averaging 27.4 years of age and 2.2 years of work experience at their current respective firms, while 98.0% had either college degrees or higher qualifications.

Measures

Collaborative HRM

The measurement of collaborative HRM concentrates on employees' abilities, motivation, and opportunities for collabation and network building. These items were adapted from three studies that explicitly use the concept of collaborative HRM (Lepak and Snell, 2002; Youndt and Snell, 2004; Lopez-Cabrales et al., 2009). At the same time, we refer to the research on SHRM with high-tech enterprises (Collins and Smith, 2006) and core knowledge employees (Kang et al., 2007) as the research objects. In order to better adapt to the Chinese context, we also conducted interviews with HR managers in IT companies to supplement the responses derived from the questionnaire survey. Finally, an initial questionnaire of 16 items was formed. We then used two methods to improve validity. On the one hand, we tested the content validity by a professional team composed of 10 management experts. The members of the expert panel responses were scored 1-4 for each item of the scale (1 for no correlation, 2 for weak correlation, 3 for strong correlation, and 4 for strong correlation). For each item, the number of experts with a score of 3 or 4 divided by the total number of experts participating in the evaluation is the content validity index (CVI) of the corresponding item. Likewise, we performed factor analysis on the initial questionnaire. Finally, we eliminated the items with CVI score below 0.8 and streamlined the scale according to the "model modification indices" provided by Mplus 8.3 (Lynn, 1986; Davis, 1992). Ultimately, a questionnaire including 11 items was formed. The measurement items includes the following: Training and development practices aimed at improving employees' cooperation and social ability; Performance, compensation, and promotion practices aimed at improving employee cooperation and social motivation; Formal work design practices aimed at enhancing opportunities for collaboration and networking among employees. A five-point Likert scale was used to measure the magnitude of all items.

Since this paper uses the scale of collaborative HRM for variable measurement, it is necessary to test the reliability and validity of this construct. Therefore, a total of 268 valid questionnaires were collected from HR managers and line managers. Exploratory factor analysis (KMO = 0.92, p < 0.001) was used to extract three factors by principal component maximum variance rotation method. The results showed that the cumulative variance contribution rate of the three factors was 73.49%, and the factor loading of each item was greater than 0.6. As shown in Table 1, the rotated component matrix shows a clear three-factor structure. The fitting index of collaborative HRM three-factor model (χ^2 /df = 2.482, RMSEA = 0.075, CFI = 0.966, TLI = 0.954, SRMR = 0.033) is shown in Table 2. Findings from

Factor	Items		Loading		Cronbach's α
CHRM-A	Providing training focusing on team- building and teamwork skills	0.76			0.86
	Providing career path opportunities for employees to move across multiple functional areas of the company	0.70			
	Sponsoring company social events for employees to become acquainted with one another	0.75			
	Conducting information-sharing meetings for employees to know more internal information about the enterprise	0.78			
CHRM-M	Promotion based on abilities to work with others		0.68		0.87
	Utilizing group-based incentives		0.79		
	Merit-based raises based on teamwork skills and team orientation		0.82		
	Monetary rewards based on the outcomes of interdependent tasks		0.82		
CHRM-O	Most of the work is performed through teamwork			0.81	0.84
	Most of work is interdependent, rather than independent of each other			0.78	
	Building cross-functional teams to complete the work			0.72	

TABLE 1 The results of exploratory factor analyses for collaborative HRM.

Extraction method, principal component analysis; Rotation method, varimax with Kaiser normalization; CHRM-A, the practices of ability dimension in collaborative HRM; CHRM-M, the practices of motivation dimension in collaborative HRM; CHRM-O, the practices of opportunity dimension in collaborative HRM.

Table 2 reveal that the fitting index of the three-factor model is significantly better than that of the two-factor model and the single-factor model. Therefore, consistent with theoretical expectations, collaborative HRM is a three-dimensional construct, which is reflected in improving employees' social ability through training and development practices, stimulating employees' social motivation through assessment, compensation and promotion practices, and providing social opportunities through formal job design. The alpha value of each dimension is greater than 0.8. Moreover, the alpha for the scale was 0.92.

Since there are criticisms regarding the use of a single response to measure company HRM practices, we then constructed the collaborative HRM index by averaging the index score of the HR manager and that of the line managers. The ICC(1) was 0.09, while the average of the $\text{Rwg}_{(j)}$ for the collaborative HRM index was 0.9, indicating that the aggregation was justified (Bliese, 2000).

Climate for cooperation

Climate for cooperation was measured by a 4-item scale adapted from Chatman and Flynn (2001). Items included "it is important to maintain harmony among employees within the organization," "employees in this organization are willing to sacrifice their self-interest for the benefit of the organization," "There is a high degree of cooperation among employees in this company," and "There is a high sense of sharing among employees in this company." The Cronbach's alpha for the scale is 0.73. We then averaged the responses of employees in each company as core knowledge employees' perceived organizational climate for cooperation. The mean $\text{Rwg}_{(j)}$ was 0.85 and the ICC(1) was 0.16, which can justify such aggregation to the organization-level in this study (Bliese, 2000).

Employee intra-organizational social ties

Following Collins and Clark (2003) and Pil and Leana (2009), employee intra-organizational social ties were measured by tie number and tie strength. The tie number was measured by the question "How many colleagues within the company do you regularly discuss expertise or exchange technical information with?." The response options range from 1 to 7 (1 represents "0," 2 represents "1," 3 represents "2," 4 represents "3," 5 represents "4," 6 represents "5," and 7 represents "6 or above"). The tie strength was measured as the average of frequency and closeness (Collins and Clark, 2003; Pil and Leana, 2009). Frequency was measured as the number of times that the employees spoke to their peers to gain expertise and related information during the last month (1 = "0-5)time(s)," 2="6-10 times," 3="11-15times," 4="16-20 times" and 5="more than 21 times"). Closeness was measured by a one-item 5-point scale "In general, how closely do you feel you discuss expertise or exchange technical information with your colleagues?" (1 = not at all close; 5 = very close).

Control variables

At the individual level, we controlled for the employees' gender, age, tenure, and education to avoid experience-related prejudice and gender effects (Snape and Redman, 2010). We also controlled for the following

TABLE 2 The results of confirmatory factor analyses for collaborative HRM.

Models	χ²	df	χ²/ df	RMSEA	CFI	TLI	SRMR
Three- factor model	101.761	41	2.482	0.075	0.966	0.954	0.033
Two-factor modelª	175.153	43	4.073	0.107	0.926	0.905	0.044
Two-factor model ^b	250.546	43	5.827	0.134	0.883	0.851	0.055
Two-factor model ^c	262.010	43	6.093	0.138	0.877	0.842	0.056
One-factor model	332.848	44	7.565	0.157	0.837	0.797	0.063

^aThe ability dimension and the opportunity dimension were combined in two-factor model. ^bThe ability dimension and the motivation dimension were combined in two-factor model. ^cThe motivation dimension and the opportunity dimension were combined in two-factor model.

TABLE 3 The results of confirmatory factor analyses for measurement model.

Models	χ²	df	χ²/ df	RMSEA	CFI	TLI	SRMR
Three- factor model	21.101	32	0.659	0.000	1.000	1.048	0.022
Two-factor modelª	94.487	35	2.700	0.069	0.835	0.759	0.097
Two-factor model ^b	136.553	39	3.501	0.083	0.729	0.646	0.127
One-factor model	279.868	42	6.664	0.125	0.340	0.198	0.180

^aClimate for cooperation and social ties were combined in a two-factor model. ^bCollaborative HRM and climate for cooperation were combined in a two-factor model .

organizational variables: firm age (years since the legal establishment), firm size (natural logarithm of incumbent full-time employees) as previous studies suggested (Takeuchi et al., 2007).

Confirmatory factor analysis and Harman's single-factor test

We conducted a series of confirmatory factor analyses (CFA) to test the measurement model specifying collaborative HRM, climate for cooperation, and employees' intra-organizational social ties as separate factors. Table 3 presents the CFA results. As shown, the hypothesized 3-factor model (χ^2 /df=0.659, RMSEA=0.000, CFI=1.000, TLI=1.048, SRMR=0.022) fits the data better than the 2-factor model^a (χ^2 /df=2.700, RMSEA=0.069, CFI=0.835, TLI=0.759, SRMR=0.097), 2-factor model^b (χ^2 /df=3.501, RMSEA=0.083, CFI=0.729, TLI=0.646, SRMR=0.127), and the 1-factor model (χ^2 /df=6.664, RMSEA=0.125, CFI=0.340, TLI=0.198, SRMR=0.180). Therefore, these results suggest that our measures' exhibit discriminant validity among these constructs, since each measure is conceptually distinct.

We next ran Harman's single-factor test using all the items at the individual level to examine the potential common method variance problem (Podsakoff et al., 2003). The result showed that the single factor accounted for approximately 27.65% of variance, which is less than the

threshold (i.e., 50%). Based on this result, we conclude that common method variance is not a serious issue in our dataset.

Analytical approach

We use SPSS 23.0 to perform the descriptive statistics and correlations, and Mplus 8.3 to test the cross-level mediation-upper mediator model.

First, we applied Kurtosis and Skewness to check the data normality following Muthén and Kaplan (1985, 1992) and Ferrando and Anguiano-Carrasco (2010), who recommend coefficients of skewness and kurtosis in a range of-1, 1. All the variables in our hypothesized model were normally distributed with Skewness ranging between-1 and 1 (Collaborative HRM = -0.033; Climate for cooperation = -0.011; Tie number=0.115; Tie strength=0.232) and Kurtosis ranging between-1 and 1 (Collaborative HRM=0.025; Climate for cooperation = -0.013; Tie number = -0.804; Tie strength = -0.505). To justify that the data is appropriate for analyzing two-level model, we began with a null model to calculate the intra-class correlation coefficient (ICC). The results provided an ICC(1) = 0.112 for employee tie number and an ICC(1) = 0.073 for employee tie strength, both of which are higher than the 0.059 recommended by Cohen (1988). The method used in the estimation of statistical outputs is robust maximum likelihood estimation, which has the advantage as less sensitive to the number of samples, while the commonly used maximum likelihood estimation is more sensitive to the number of samples and missing data.

To test the significance of the indirect effects, A Monte Carlo simulation with 20,000 replications was conducted to test the 95% biascorrected confidence interval (CI), using the web estimator provided by Selig and Preacher (2008). This method has been suggested to determine indirect effects in multilevel models (Preacher and Selig, 2012). If the 95% CI does not include zero, we can conclude that the indirect effect is significant. Based on Hu and Bentler (1999) recommendations, means of the root-mean-square error of approximation (RMSEA), the Tucker-Lewis Index (TLI), and the comparative fit index (CFI), and the standardized root mean square residual (SRMR) were employed to assess model fit. The following cut-off values were used: RMSEA values below 0.08, TLI values higher than 0.8, CFI values higher than 0.8, and SRMR values below 0.08.

Results

Descriptive statistics and correlations

The mean, standard deviation, and correlation coefficient matrix of the variables are shown in Table 4.

Hypothesis tests

The results of hypothesis test are shown in Table 5. In Model 1, we estimated the independent variable collaborative HRM's direct effect on the dependent variable (i.e., tie number and tie strength), with individual-level control variables and organizational-level control variables. The results show collaborative HRM has a positive effect on tie number (b = 0.574, p < 0.05) and tie strength (b = 0.406, p < 0.01). Therefore, Hypothesis 1 is supported.

In Model 2, we estimated the direct effect of the independent variable (i.e., collaborative HRM) on the mediating variable (i.e., climate

TABLE 4 Descriptive statistics, means, standard deviations, and correlations.

Variables	М	SD	1	2	3	4	5	6
Individual level								
(N = 363)								
1. Age	27.43	4.79						
2. Gender	1.45	0.49	0.79					
3. Edu	2.89	0.58	0.27	0.49				
4. Work tenure	2.20	1.93	0.49**	0.77	-0.13*			
5. CC	3.66	0.74	-0.76	0.88	-0.07	0.01		
6. TN	4.66	1.54	-0.18	-0.02	0.07	-0.03	0.31**	
7. TS	2.96	0.94	-0.13*	-0.09	0.01	-0.06	0.26**	0.49**
Firm level $(N = 64)$								
1. Firm age	11.84	7.12						
2. Firm size	5.94	1.29	0.41**					
3. CHRM	3.40	0.47	-0.14	-0.04				
4. CC	3.67	0.44	-0.31	0.12	0.31*			

CHRM, collaborative HRM; CC, climate for cooperation; TN, tie number; TS, tie strength. *p < 0.05, **p < 0.01.

TABLE 5 The model results.

Variables	Mo	del 1	Model 2	Мос	del 3		Model 4		
	TN	TS	CC	TN	TS	CC	TN	TS	
Intercept	1.679	2.105**	2.350***	0.208	2.227***	5.662***	-0.454	1.43*	
Individual level									
Age	0.004	-0.018		-0.003	-0.024		0.001	-0.019	
Gender	-0.097	-0.187	-	-0.072	-0.166		-0.092	-0.189	
edu	0.159	0.033		0.137	0.007		0.144	0.027	
Work tenure	-0.011	0.006	-	-0.008	0.020		-0.008	0.007	
Organizational level									
Firm age	-0.003	0.004	0.000	-0.003	0.002	-0.049	-0.001	0.005	
Firm size	0.108	0.011	0.025	0.078	-0.003	0.150	0.082	0.003	
CHRM	0.574*	0.406**	0.338***			0.307*	0.242	0.296*	
CC				1.048***	0.419**		0.956***	0.309*	
R ²	0.450	0.567	0.135	0.884	0.590	0.114	0.896	0.809	

n = 363 for individual-level variables. n = 64 for organization-level variables. CHRM, collaborative HRM; CC, climate for cooperation; TN, tie number; TS, tie strength. The coefficient is non-standard coefficient; R^2 calculated by Mplus 8.3 according to the normalization coefficient. *p < 0.05, **p < 0.01, ***p < 0.001.

for cooperation). The regression results show that a collaborative HRM system is positively related to the climate for cooperation (b = 0.338, p < 0.001).

In Model 3, we examined the effect of the mediating variable (i.e., climate for cooperation) on the dependent variable (i.e., tie number and tie strength). Our results indicate that the climate for cooperation was significantly associated with tie number (b = 1.048, p < 0.001) and tie strength (b = 0.419, p < 0.01).

In Model 4, we examined the mediating role of climate for cooperation. When we put the independent variable and the mediating variable together in the regression model, the results show that the mediating variable (i.e., climate for cooperation) has a significant positive relationship with the number (b = 0.956, p < 0.001) and the strength (b = 0.309, p < 0.05), whereas the

relationship between collaborative HRM and tie number decreased from 0.574 (p < 0.05) to 0.242 (p > 0.05), and the relationship between collaborative HRM and tie strength decreased from 0.406 (p < 0.01) to 0.296 (p < 0.05). Results showed that the indirect effects between collaborative HRM and tie number (indirect effect = 0.271, 95%CI = [0.0393, 0.5596]) and tie strength (indirect effect = 0.087, 95%CI = [0.0041, 0.2121]) through climate for cooperation were significant. The overall results of the proposed model indicated an acceptable level of model fit (RMSEA = 0.000, TLI = 1.000, CFI = 1.000, SRMR within = 0.002, SRMR between = 0.023). Consequently, we concluded that a climate for cooperation mediates the relationship between collaborative HRM, and the number and strength of employees' intra-organizational social ties. Thus Hypothesis 2 is supported (see Figure 2).



TABLE 6 Comparison of total effects between different dimensions in CHRM.

Variables		R ²		Total effect			
	ΤN	TS	СС	ΤN	TS	СС	
CHRM-A	0.303	0.400	0.093	0.389*	0.294**	0.238***	
CHRM-M	0.355	0.526	0.124	0.405	0.325**	0.270***	
CHRM-AM	0.363	0.527	0.122	0.452*	0.352**	0.288***	
CHRM-AMO	0.450	0.567	0.135	0.574*	0.406**	0.338***	

CHRM, collaborative HRM; CHRM-A, the practices of ability dimension in CHRM; CHRM-M, the practices of motivation dimension in CHRM; CHRM-AM, the combination of the practice of ability dimension and the practice of motivation dimension in CHRM; CC, climate for cooperation; TN, tie number; TS, tie strength. the coefficient is non-standard coefficient; R² calculated by Mplus 8.3 according to the normalization coefficient. *p <0.05, **p <0.01, ***p <0.001.

Additional tests

In previous studies, the collaborative HRM construct has been used to emphasize the dimensions of ability and motivation (Lopez-Cabrales et al., 2009), however, this study integrates the opportunity dimension into this construct. Hence, we compared both the total and indirect effects of collaborative HRM as proposed in this study, and the traditional collaborative HRM model which includes only the ability and motivation dimensions. The ensuing comparison of these total effects are shown in Table 6, and the comparison of indirect effects to the adjusted model are shown in Table 7. The results show that the effectiveness of collaborative HRM with three dimensions of ability, motivation and opportunity is better than that with only two dimensions of ability and motivation.

Discussion

Theoretical implications

The objective of this study is to examine the cross-level relationship between collaborative HRM and employee's intra-organizational social ties,

TABLE 7 Comparison of indirect effects between different dimensions in CHRM.

Variables		R ²		Indirect effect		90% percent		
	ΤN	TS	СС	ΤN	TS	ΤN	TS	
CHRM-A	0.891	0.762	0.098	0.225 (0.048)	0.078 (0.086)	[0.038, 0.413]	[0.003, 0.153]	
CHRM-M	0.878	0.775	0.083 0.099	0.192 (0.081)	0.063 (0.136)	[0.011, 0.372]	[-0.006, 0.131]	
CHRM-AM	0.885	0.797		0.232 (0.052)	0.076 (0.101)	[0.036, 0.429]	[0.000, 0.152]	
CHRM- AMO	0.896	0.809	0.114	0.271 (0.035)	0.087 (0.091)	[0.059, 0.482]	[0.002, 0.172]	

CHRM, collaborative HRM; CHRM-A, the practices of ability dimension in CHRM; CHRM-M, the practices of motivation dimension in CHRM; CHRM-AM, the combination of the practice of ability dimension and motivation dimension in CHRM; TN, tie number; TS, tie strength. The coefficient is non-standard coefficient; R^2 calculated by Mplus 8.3 according to the normalization coefficient. () = value of p appear in parentheses.

as well as the mediating role of the climate for cooperation in this relationship. This study has several implications for both theory and practice.

First, our study revealed that organizational level collaborative HRM and climate for cooperation positively influence employees' intraorganizational social ties, which contributes to the social network literature since few studies have investigated how organizational contexts affect employee social networks in a cross-level way. Especially in high-tech enterprises, social ties within an organization play a key role in knowledge transfer and sharing (Hu, 2008; Zhou et al., 2009), thus organizations need to guide the cooperation and exchange among employees through wellarticulated policies such as HRM (Patel et al., 2013; Garaus et al., 2016). Although this phenomenon has received increasingly attention from scholars in recent years (Methot et al., 2018; Soltis et al., 2018), the empirical research about cross-level effects of organizational context factor on employees' social ties is still rare. Thus, this paper expands the antecedent variables in the field of social networks.

Second, the finding highlights the importance of collaborative HRM practices in knowledge intensive firms, and reveals an important

path for collaborative HRM to exert its effectiveness by incorporating the organizational climate. Since the social interactions and social networks play important roles in knowledge sharing and creation (Reagans and McEvily, 2003; Hansen et al., 2005; Chen and Huang, 2007), HRM in organizations need promote the social relationships among employees. However, a large number of previous SHRM studies have focused on high-performance work systems, with limited investigations of other forms of HRM. Our study found that collaborative HRM practices have a significant cross-level influence on employee intra-organizational social ties, which enrich our understanding of the collaborative HRM and call for a deeper and more theoretical interpretation of SHRM research in the knowledge economy. Besides, our study found that organizational climate for cooperation play a mediating role, which is consistent with the findings of previous studies that have taken the social climate as a mediation variable of SHRM (Prieto and Santana, 2012; Cooper et al., 2019).

Third, using the AMO model, our study improves the construct and measurement of collaborative HRM, thus expanding the configuration of SHRM. In the SHRM field, different configurations and specific-target orientations are becoming more and more important (Shen and Benson, 2014; Chuang et al., 2016; Lee et al., 2019). Collaborative HRM is an important configuration of SHRM, which is characterized by investing in interpersonal relationships. However, in the process of developing this construct, previous studies did not adequately use relevant theories to define and examine the measure structure of this important construct, and did not conduct a CFA test. Compared with the previous research (Lopez-Cabrales et al., 2009), we proposed a clearer abilitymotivation-opportunity three-dimensional structure of collaborative HRM. Through factor analysis, hypothesis testing and additional testing, we verified the dimensions of the construct, and its effect on employees' social ties, thus enriching our understanding of collaborative HRM, as well as expanding the configuration of SHRM from the perspective of social network.

Practical implications

This study has important implications for management practices. This is because the competitive advantage of high-tech firms are derived from knowledge sharing and creation. Interestingly, collaborative HRM can create a cooperative climate where employees collaborate with their colleagues, and also expand their social ties within an organization, consequently promoting knowledge sharing and creation. This implies that for IT firms, it is risky to simply invest in employees' human capitals. Given the fact that there is fierce talent competition and a high rate of turnover in the IT industry, we suggest that this may impose an organization to a high investment risk. Comparatively, rather than investing in individual employees, collaborative HRM emphasizes investment in employees' social networks, which are embedded within organizations and cannot be easily dissolved when a focal employee leaves a company. Thus, it provides hightech firms with both a useful tool to promote knowledge exchange, and an opportunity to enhance the human capital skillsets of their employees. Therefore, designing and implementing collaborative HRM systems in an organization is essential for IT firms, in order to gain competitive advantage in a knowledge-driven economy.

Our study has the following limitations. First, because studies on collaborative HRM are relatively few and sketchy, the items in our questionnaire may not completely cover every aspect of a collaborative HRM system. Future studies should therefore explore the composition of a collaborative HRM system in greater details. Second, because the goal of this study is to investigate how organizational factors influence employee social ties, we measured the collaborative HRM at the organizational level using the managers reports rather than employees' perceptions of HR practices. However, previous studies have revealed that employees' perceptions of HR practices can significantly influence their work attitudes and behaviors, and may also mediate the relationships between organization intended HR practices and employee outcomes (Wang et al., 2020). Thus future research can incorporate employee percpetions of HR practices by investigating how organizations can improve employees' perceptions of a collaborative HRM system, and the conditions under which employees' perceptions of collaborative HR practices can mediate the relationships between organizational level collaborative HRM and individual level social ties. Third, one assumption of this study is that employee social ties are positively related to knowledge sharing and creation, and ultimately employee performance. Although many empirical studies have provided the related evidence, our empirical investigation did not include knowledge sharing and creation, and employee performance. Future studies may examine the causal chain linking collaborative HRM and employee performance as a complete model. Fourth, though we had several characteristic variables of employees and organizations as the controls, we did not consider the influences of personality characteristics on individual social networks. Past research has demonstrated that personal characterstics such as proactive personality (Thompson, 2005), selfmonitors (Sasovova et al., 2010) and cooperative orientation (Chen X. P. et al., 2011) can influence personal social network formation. For future research, on the one hand, these personality traits can be included as control variables, so as to study the impacts of collaborative HRM on employees' social ties in a more accurate way. On the other hand, these personality traits may be regarded as boundary condition. For examples, employees with high proactive personality or cooperative orientation may hold more intrinsic motivations to expand their social interactions, which may weaken the impacts of collaborative HRM on employees' social ties. Thus, future research can include employee personality variables in the research model to better understand the relationships between HR practices and social ties from a more employee-focused perspective.

Conclusion

In closing, the primary goal of this study was to examine how the organizational-level collaborative HRM systems affect individual-level intra-organizational social ties. Based on the analysis of the dataset comprising 363 knowledge employees from 64 high-tech firms in China, the results showed that collaborative HRM had significant positive effects on the number and strength of individual-level intra-organizational social ties, and the climate for cooperation mediates these relationships. Our study not only advances the knowledge concerning how organizational contexts affect employee social networks in a

cross-level way, but also reveals an important path for collaborative HRM to exert its influences by considering the organizational climate.

Data availability statement

The raw data supporting the conclusions of this article will be made available by the authors, without undue reservation.

Ethics statement

The studies involving human participants were reviewed and approved by Academic Ethics Committee of Renmin University of China. The patients/participants provided their written informed consent to participate in this study. Written informed consent was obtained from the individual(s) for the publication of any potentially identifiable images or data included in this article.

Author contributions

ZS and XL: conceptualization, methodology, and writing-original draft preparation. MZ and XL: data curation, software, writing, reviewing, and editing. MZ, YY, and WS: visualization, investigation,

References

Appelbaum, E., Bailey, T. R., Berg, P. B., and Kalleberg, A. L. (2000). *Manufacturing advantage: Why high-performance work systems pay off.* Ithaca, N.Y.: Cornell University Press.

Bliese, P. D. (2000). Within-group agreement, non-independence, and reliability: Implications for data aggregation and analysis. San Francisco, CA: Jossey-Bass.

Borgatti, S.P., Everett, M.G., and Johnson, J.C. (2013). *Analyzing social networks*. Thousand Oaks, CA: Sage.

Bouty, I. (2000). Interpersonal and interaction influences on informal resource exchanges between R&D researchers across organizational boundaries. *Acad. Manag. J.* 43, 50–65. doi: 10.5465/1556385

Bowen, D. E., and Ostroff, C. (2004). Understanding HRM-firm performance linkages: the role of the "strength" of the HRM system. *Acad. Manag. Rev.* 29, 203–221. doi: 10.5465/amr.2004.12736076

Brass, D. J., Galaskiewicz, J., Greve, H. R., and Tsai, W. (2004). Taking stock of networks and organizations: a multilevel perspective. *Acad. Manag. J.* 47, 795–817. doi: 10.5465/20159624

Burt, R. S. (1997). A note on social capital and network content. Soc. Networks 19, 355–373. doi: 10.1016/S0378-8733(97)00003-8

Chatman, J. A., and Flynn, F. J. (2001). The influence of demographic heterogeneity on the emergence and consequences of cooperative norms in work teams. *Acad. Manag. J.* 44, 956–974. doi: 10.5465/3069440

Chen, C. J., and Huang, J. W. (2007). How organizational climate and structure affect knowledge management—the social interaction perspective. *Int. J. Inf. Manag.* 27, 104–118. doi: 10.1016/j.ijinfomgt.2006.11.001

Chen, X. P., Xie, X., and Chang, S. (2011). Cooperative and competitive orientation among Chinese people: scale development and validation. *Manag. Organ. Rev.* 7, 353–379. doi: 10.1111/j.1740-8784.2011.00215.x

Chen, Y. Y., Zhang, Y., and Fey, C. F. (2011). When collaborative HR practices may not work well: the moderating role of social capital in the Chinese life insurance industry. *Int. J. Hum. Resour. Manag.* 22, 433–456. doi: 10.1080/09585192.2011. 540164

Chuang, C. H., Jackson, S. E., and Jiang, Y. (2016). Can knowledge-intensive teamwork be managed? Examining the roles of HRM systems, leadership, and tacit knowledge. *J. Manag.* 42, 524–554. doi: 10.1177/0149206313478189

Cohen, J. (1988). *Statistical power analysis for the behavioral sciences (2nd Ed.)*. Hillsdale, NJ: Lawrence Erlbaum Associates.

Colakoglu, S., Chung, Y., and Ceylan, C. (2021). Collaboration-based HR systems and innovative work behaviors: the role of information exchange and HR system strength. *Eur. Manag. J.* 40, 518–531. doi: 10.1016/j.emj.2021.07.011

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Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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Coleman, J. S. (1988). Social capital in the Creation of human capital. Am. J. Sociol. 94, S95–S120. doi: 10.1086/228943

Collins, C. J., and Clark, K. D. (2003). Strategic human resource practices, top management team social networks, and firm performance: the role of human resource practices in creating organizational competitive advantage. *Acad. Manag. J.* 46, 740–751. doi: 10.5465/30040665

Collins, C. J., and Smith, K. G. (2006). Knowledge exchange and combination: the role of human resource practices in the performance of high-technology firms. *Acad. Manag. J.* 49, 544–560. doi: 10.5465/amj.2006.21794671

Combs, J., Liu, Y., Hall, A., and Ketchen, D. (2006). How much do high-performance work practices matter? A meta-analysis of their effects on organizational performance. *Pers. Psychol.* 59, 501–528. doi: 10.1111/j.1744-6570.2006.00045.x

Cooper, B., Wang, J., Bartram, T., and Cooke, F. L. (2019). Well-being-oriented human resource management practices and employee performance in the Chinese banking sector: the role of social climate and resilience. *Hum. Resour. Manag.* 58, 85–97. doi: 10.1002/hrm.21934

Davenport, T. H., and Prusak, L. (1998). Working knowledge: How organizations manage what they know. Boston, Massachusetts: Harvard Business Press.

Davis, L. L. (1992). Instrument review: getting the most from a panel of experts. Appl. Nurs. Res. 5, 194–197. doi: 10.1016/S0897-1897(05)80008-4

Dean, J. W. Jr., and Brass, D. J. (1985). Social interaction and the perception of job characteristics in an organization. *Hum. Relat.* 38, 571–582. doi: 10.1177/001872678503800605

Evans, W. R., and Davis, W. D. (2005). High-performance work systems and organizational performance: the mediating role of internal social structure. *J. Manag.* 31, 758–775. doi: 10.1177/0149206305279370

Ferrando, P. J., and Anguiano-Carrasco, C. (2010). El análisis factorial como técnica de investigación en psicología [Factor analysis as a technique in psychological research]. *Papeles del Psicól.* 31, 18–33. https://www.redalyc.org/articulo. oa?id=77812441003

Ferris, G. R., Arthur, M. M., Berkson, H. M., Kaplan, D. M., Harrell-Cook, G., and Frink, D. D. (1998). Toward a social context theory of the human resource managementorganization effectiveness relationship. *Hum. Resour. Manag. Rev.* 8, 235–264. doi: 10.1016/ S1053-4822(98)90004-3

Garaus, C., Güttel, W. H., Konlechner, S., Koprax, I., Lackner, H., Link, K., et al. (2016). Bridging knowledge in ambidextrous HRM systems: empirical evidence from hidden champions. *Int. J. Hum. Resour. Manag.* 27, 355–381. doi: 10.1080/09585192.2015.1045007

Guimera, R., Uzzi, B., Spiro, J., and Amaral, L. A. N. (2005). Team assembly mechanisms determine collaboration network structure and team performance. *Science* 308, 697–702. doi: 10.1126/science.1106340

Hansen, M. T., Mors, M. L., and Løvås, B. (2005). Knowledge sharing in organizations: multiple networks, multiple phases. *Acad. Manag. J.* 48, 776–793. doi: 10.5465/amj.2005.18803922

Heffernan, M., and Dundon, T. (2016). Cross-level effects of high-performance work systems (HPWS) and employee well-being: the mediating effect of organisational justice. *Hum. Resour. Manag. J.* 26, 211–231. doi: 10.1111/1748-8583.12095

Hirst, G., Van Knippenberg, D., Zhou, J., Quintane, E., and Zhu, C. (2015). Heard it through the grapevine: indirect networks and employee creativity. *J. Appl. Psychol.* 100, 567–574. doi: 10.1037/a0038333

Hollenbeck, J. R., and Jamieson, B. B. (2015). Human capital, social capital, and social network analysis: implications for strategic human resource management. *Acad. Manag. Perspect.* 29, 370–385. doi: 10.5465/amp.2014.0140

Hu, T. S. (2008). Interaction among high-tech talent and its impact on innovation performance: a comparison of Taiwanese science parks at different stages of development. *Eur. Plan. Stud.* 16, 163–187. doi: 10.1080/09654310701814462

Hu, L. T., and Bentler, P. M. (1999). Cutoff criteria for fit indexes in covariance structure analysis: conventional criteria versus new alternatives. *Struct. Equ. Model. Multidiscip. J.* 6, 1–55. doi: 10.1080/10705519909540118

Huang, Y., Ma, Z., and Meng, Y. (2018). High-performance work systems and employee engagement: empirical evidence from China. *Asia Pac. J. Hum. Resour.* 56, 341–359. doi: 10.1111/1744-7941.12140

Huselid, M. A. (1995). The impact of human resource management practices on turnover, productivity, and corporate financial performance. *Acad. Manag. J.* 38, 635–672. doi: 10.5465/256741

Ingram, P., and Roberts, P. W. (2000). Friendships among competitors in the Sydney hotel industry. *Am. J. Sociol.* 106, 387–423. doi: 10.1086/316965

Jiang, K., Lepak, D. P., Hu, J., and Baer, J. C. (2012). How does human resource management influence organizational outcomes? A meta-analytic investigation of mediating mechanisms. *Acad. Manag. J.* 55, 1264–1294. doi: 10.5465/amj.2011.0088

Kale, P., Singh, H., and Perlmutter, H. (2000). Learning and protection of proprietary assets in strategic alliances: building relational capital. *Strateg. Manag. J.* 21, 217–237. doi: 10.1002/(SICI)1097-0266(200003)21:3<217::AID-SMJ95>3.0.CO;2-Y

Kang, S. C., Morris, S. S., and Snell, S. A. (2007). Relational archetypes, organizational learning, and value creation: extending the human resource architecture. *Acad. Manag. Rev.* 32, 236–256. doi: 10.5465/amr.2007.23464060

Kaše, R., Paauwe, J., and Zupan, N. (2009). Hr practices, interpersonal relations, and intrafirm knowledge transfer in knowledge-intensive firms: a social network perspective. *Hum. Resour. Manag.* 48, 615–639. doi: 10.1002/hrm.20301

Kehoe, R. R., and Collins, C. J. (2017). Human resource management and unit performance in knowledge-intensive work. J. Appl. Psychol. 102, 1222–1236. doi: 10.1037/apl0000216

Kehoe, R. R., and Wright, P. M. (2013). The impact of high-performance human resource practices on employees' attitudes and behaviors. *J. Manag.* 39, 366–391. doi: 10.1177/0149206310365901

Kim, S., Suh, E., and Jun, Y. (2011). Building a knowledge brokering system using social network analysis: a case study of the Korean financial industry. *Expert Syst. Appl.* 38, 14633–14649. doi: 10.1016/j.eswa.2011.05.019

Leana, C. R. III, and Van Buren, H. J. (1999). Organizational social capital and employment practices. Acad. Manag. Rev. 24, 538–555. doi: 10.5465/amr.1999.2202136

Leana, C. R., and Pil, F. K. (2006). Social capital and organizational performance: evidence from urban public schools. *Organ. Sci.* 17, 353–366. doi: 10.1287/orsc.1060.0191

Lee, H. W., Pak, J., Kim, S., and Li, L. Z. (2019). Effects of human resource management systems on employee proactivity and group innovation. *J. Manag.* 45, 819–846. doi: 10.1177/0149206316680029

Lengnick-Hall, M. L., and Lengnick-Hall, C. A. (2006). "International human resource management and social network/social capital theory" in *The handbook of research in international human resource management*. eds. G. K. Stahl and I. Björkman (Cheltenham, United Kingdom: Edward Elgar), 475. doi: 10.4337/9781845428235.00034

Lepak, D. P., Liao, H., Chung, Y., and Harden, E. E. (2006). A conceptual review of human resource management systems in strategic human resource management research. *Res. Pers. Hum. Resour. Manag.* 25, 217–271. doi: 10.1016/S0742-7301(06)25006-0

Lepak, D. P., and Snell, S. A. (1999). The human resource architecture: toward a theory of human capital allocation and development. *Acad. Manag. Rev.* 24, 31–48. doi: 10.5465/amr.1999.1580439

Lepak, D. P., and Snell, S. A. (2002). Examining the human resource architecture: the relationships among human capital, employment, and human resource configurations. *J. Manag.* 28, 517–543. doi: 10.1177/014920630202800403

Liebeskind, J. P., Oliver, A. L., Zucker, L., and Brewer, M. (1996). Social networks, learning, and flexibility: sourcing scientific knowledge in new biotechnology firms. *Organ. Sci.* 7, 428–443. doi: 10.1287/orsc.7.4.428

Liu, M. S., and Liu, N. C. (2008). Sources of knowledge acquisition and patterns of knowledge-sharing behaviors—an empirical study of Taiwanese high-tech firms. *Int. J. Inf. Manag.* 28, 423–432. doi: 10.1016/j.ijinfomgt.2008.01.005

Lopez-Cabrales, A., Pérez-Luño, A., and Cabrera, R. V. (2009). Knowledge as a mediator between HRM practices and innovative activity. *Hum. Resour. Manag.* 48, 485–503. doi: 10.1002/hrm.20295

Louis, M. R. (1980). Surprise and sense making: what newcomers experience in entering unfamiliar organizational settings. *Adm. Sci. Q.* 25:2453. doi: 10.2307/2392453

Lynn, M. R. (1986). Determination and quantification of content validity. Nurs. Res. 35, 382???386-382???385. doi: 10.1097/00006199-198611000-00017

Methot, J. R., Rosado-Solomon, E. H., and Allen, D. G. (2018). The network architecture of human captial: a relational identity perspective. *Acad. Manag. Rev.* 43, 723–748. doi: 10.5465/amr.2016.0338

Müller-Prothmann, T. (2007). "Social network analysis: a practical method to improve knowledge sharing" in *The hands-on knowledge co-creation and sharing: Practical methods and techniques*. eds. A. S. Kazi, L. Wohlfahrt and P. Wolf (Stuttgart, Germany: Knowledge Board), 219–233.

Muthén, B., and Kaplan, D. (1985). A comparison of some methodologies for the factor analysis of non-normal Likert variables. *Br. J. Math. Stat. Psychol.* 38, 171–189. doi: 10.1111/j.2044-8317.1985.tb00832.x

Muthén, B., and Kaplan, D. (1992). A comparison of some methodologies for the factor analysis of non-normal Likert variables: a note on the size of the model. *Br. J. Math. Stat. Psychol.* 45, 19–30. doi: 10.1111/j.2044-8317.1992.tb00975.x

Nahapiet, J., and Ghoshal, S. (1998). Social capital, intellectual capital, and the organizational advantage. Acad. Manag. Rev. 23, 242–266. doi: 10.5465/amr.1998.533225

Nieves, J., Quintana, A., and Osorio, J. (2016). Organizational knowledge and collaborative human resource practices as determinants of innovation. *Knowl. Manag. Res. Pract.* 14, 237–245. doi: 10.1057/kmrp.2014.26

Oh, H., Labianca, G., and Chung, M. H. (2006). A multilevel model of group social capital. Acad. Manag. Rev. 31, 569-582. doi: 10.5465/amr.2006.21318918

Osman-Gani, A., and Rockstuhl, T. (2009). Antecedents and consequences of social network characteristics for expatriate adjustment and performance in overseas assignments: implications for HRD. *Hum. Resour. Dev. Rev.* 7, 32–57. doi: 10.1177/153 4484307312182

Pahos, N., and Galanaki, E. (2022). Performance effects of high performance work systems on committed, long-term employees: a multilevel study. *Front. Psychol.* 13:825397. doi: 10.3389/fpsyg.2022.825397

Patel, P. C., Messersmith, J. G., and Lepak, D. P. (2013). Walking the tightrope: an assessment of the relationship between high-performance work systems and organizational ambidexterity. *Acad. Manag. J.* 56, 1420–1442. doi: 10.5465/amj.2011.0255

Pil, F. K., and Leana, C. (2009). Applying organizational research to public school reform: the effects of teacher human and social capital on student performance. *Acad. Manag. J.* 52, 1101–1124. doi: 10.5465/amj.2009.47084647

Podsakoff, P. M., MacKenzie, S. B., Lee, J. Y., and Podsakoff, N. P. (2003). Common method biases in behavioral research: a critical review of the literature and recommended remedies. *J. Appl. Psychol.* 88, 879–903. doi: 10.1037/0021-9010.88.5.879

Preacher, K. J., and Selig, J. P. (2012). Advantages of Monte Carlo confidence intervals for indirect effects. Confidence intervals for indirect effects. *Commun. Methods Meas.* 6, 77–98. doi: 10.1080/19312458.2012.679848

Prieto, I. M., and Santana, P. P. M. (2012). Building ambidexterity: the role of human resource practices in the performance of firms from Spain. *Hum. Resour. Manag.* 51, 189–211. doi: 10.1002/hrm.21463

Reagans, R., and McEvily, B. (2003). Network structure and knowledge transfer: the effects of cohesion and range. *Adm. Sci. Q.* 48, 240–267. doi: 10.2307/3556658

Rivera, M. T., Soderstrom, S. B., and Uzzi, B. (2010). Dynamics of dyads in social networks: assortative, relational, and proximity mechanisms. *Annu. Rev. Sociol.* 36, 91–115. doi: 10.1146/annurev.soc.34.040507.134743

Sasovova, Z., Mehra, A., Borgatti, S. P., and Schippers, M. C. (2010). Network churn: the effects of self-monitoring personality on brokerage dynamics. *Adm. Sci. Q.* 55, 639–670. doi: 10.2189/asqu.2010.55.4.639

Schneider, B., White, S. S., and Paul, M. C. (1998). Linking service climate and customer perceptions of service quality: tests of a causal model. *J. Appl. Psychol.* 83, 150–163. doi: 10.1037/0021-9010.83.2.150

Seibert, S. E., Kraimer, M. L., and Liden, R. C. (2001). A social capital theory of career success. Acad. Manag. J. 44, 219–237. doi: 10.5465/3069452

Selig, J. P., and Preacher, K. J. (2008). Monte Carlo method for assessing mediation: An interactive tool for creating confidence intervals for indirect effects [computer software]. Available online at: http://quantpsy.org/

Senaratne, S., Rodrigo, M. N. N., Jin, X., and Perera, S. (2021). Current trends and future directions in knowledge management in construction research using social network analysis. *Buildings* 11:599. doi: 10.3390/buildings11120599

Shen, J., and Benson, J. (2014). When CSR is a social norm: how socially responsible human resource management affects employee work behavior. *J. Manag.* 42, 1723–1746. doi: 10.1177/0149206314522300

Snape, E., and Redman, T. (2010). HRM practices, organizational citizenship behaviour, and performance: a multi-level analysis. *J. Manag. Stud.* 47, 1219–1247. doi: 10.1111/j.1467-6486.2009.00911.x

Soltis, S. M., Brass, D. J., and Lepak, D. P. (2018). Social resource management: integrating social network theory and human resource management. *Acad. Manag. Ann.* 12, 537–573. doi: 10.5465/annals.2016.0094

Sosa, M. E. (2011). Where do creative interactions come from? The role of tie content and social networks. Organ. Sci. 22, 1–21. doi: 10.1287/orsc.1090.0519

Takeuchi, R., Chen, G., and Lepak, D. P. (2009). Through the looking glass of a social system: cross-level effects of high-performance work systems on employees' attitudes. *Pers. Psychol.* 62, 1–29. doi: 10.1111/j.1744-6570.2008.01127.x

Takeuchi, R., Lepak, D. P., Wang, H., and Takeuchi, K. (2007). An empirical examination of the mechanisms mediating between high-performance work systems and the performance of Japanese organizations. *J. Appl. Psychol.* 92, 1069–1083. doi: 10.1037/0021-9010.92.4.1069

Thompson, J. A. (2005). Proactive personality and job performance: a social capital perspective. J. Appl. Psychol. 90, 1011–1017. doi: 10.1037/0021-9010.90.5.1011

Tsai, W., and Ghoshal, S. (1998). Social capital and value creation: the role of intrafirm networks. *Acad. Manag. J.* 41, 464–476. doi: 10.5465/257085

Wang, Y., Kim, S., Rafferty, A., and Sanders, K. (2020). Employee perceptions of HR practices: a critical review and future directions. *Int. J. Hum. Resour. Manag.* 31, 128–173. doi: 10.1080/09585192.2019.1674360

Youndt, M. A., and Snell, S. A. (2004). Human resource configurations, intellectual capital, and organizational performance. *J. Manag. Issues* 16, 337–360. doi: 10.jstor.org/stable/40604485

Zhang, J., Akhtar, M. N., Bal, P. M., Zhang, Y., and Talat, U. (2018). How do highperformance work systems affect individual outcomes: a multilevel perspective. *Front. Psychol.* 9:586. doi: 10.3389/fpsyg.2018.00586

Zhang, Z., Zyphur, M. J., and Preacher, K. J. (2009). Testing multilevel mediation using hierarchical linear models: problems and solutions. *Organ. Res. Methods* 12, 695–719. doi: 10.1177/1094428108327450

Zhou, J., Shin, S. J., Brass, D. J., Choi, J., and Zhang, Z.-X. (2009). Social networks, personal values, and creativity: evidence for curvilinear and interaction effects. *J. Appl. Psychol.* 94, 1544–1552. doi: 10.1037/a0016285

Zhou, Y., Zhang, Y., and Liu, J. (2012). A hybridism model of differentiated human resource management effectiveness in Chinese context. *Hum. Resour. Manag. Rev.* 22, 208–219. doi: 10.1016/j.hrmr.2012.01.003