

1 Seasonal variation in mortality and the role of temperature: a multi-country multi-city study

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46 **Abstract**

47 **Background:** Although seasonal variations in mortality have been recognised for millennia,
48 the role of temperature remains unclear. We aimed to assess seasonal variation in mortality and
49 to examine the contribution of temperature.

50 **Methods:** We compiled daily data on all-cause, cardiovascular, and respiratory mortality,
51 temperature, and indicators on location-specific characteristics from 719 locations in tropical,
52 dry, temperate and continental climate zones. We fitted time-series regression models to
53 estimate the amplitude of seasonal variation in mortality on a daily basis, defined as the peak-
54 to-trough ratio (PTR) of maximum mortality estimates to minimum mortality estimates at day-
55 of-year. Meta-analysis was used to summarise location-specific estimates for each climate
56 zone. We estimated PTR with and without temperature adjustment, with the differences
57 representing the seasonal effect attributable to temperature. We also evaluated the effect of
58 location-specific characteristics on PTR across locations by using meta-regression models.

59 **Results:** Seasonality estimates and responses to temperature adjustment varied across
60 locations. Unadjusted-PTR for all-cause mortality was 1.05 (95% confidence interval (CI):
61 1.00–1.11) in the tropical zone and 1.23 (95% CI: 1.20–1.25) in the temperate zone; adjusting
62 for temperature reduced the estimates to 1.02 (95% CI: 0.95–1.09) and 1.10 (95% CI: 1.07–
63 1.12), respectively. Furthermore, unadjusted-PTR was positively associated with average mean
64 temperature.

65 **Conclusions:** This study suggests that seasonality of mortality is importantly driven by
66 temperature, most evidently in temperate/continental climate zones, and that warmer locations
67 show stronger seasonal variations in mortality, which is related to a stronger effect of
68 temperature.

69

70 **Key words:** Seasonality, mortality, temperature

71 **Key messages:**

- 72 • To our knowledge, this is by far the largest study on seasonality of mortality by including
73 719 locations from 34 countries in tropical, dry, temperate and continental climate zones.
- 74 • Our study provides evidence that the generally higher mortality in cold seasons than in
75 warm seasons is considerably explained by temperature, and this pattern is most evident in
76 temperate and continental climate zones.
- 77 • Seasonality estimates and responses to temperature adjustment varied across locations, and
78 locations characterised by warm climate experienced larger seasonal variations in mortality,
79 which was related to stronger effect of temperature.
- 80 • Our investigation of this long-known complex phenomenon provides important evidence
81 for understanding the epidemiology and ecology of seasonal variation in mortality and the
82 role of temperature.
- 83 • Our findings also provide a basis for developing hypotheses about the potential impact of
84 climate change on seasonality of mortality for future investigations.

85 **Introduction**

86 Seasonal variation in mortality as a broad phenomenon has been recognised since Hippocrates.¹
87 During certain times of the year, mortality increases substantially, which consequently
88 increases demand for healthcare services and may exert intense pressure on healthcare
89 systems.^{2–10} The most plausible underlying mechanisms include seasonal fluctuations in
90 environment, human behaviors, and infectious diseases.

91 Although consensus exists among researchers that ambient temperature is a key driver, the
92 extent to which temperature is actually the proximal cause of seasonal variation in mortality is
93 a matter of long-standing debate.^{11,12} Few studies have attempted to address this topic.⁶ More
94 importantly, previous investigations focused on a small number of locations within a limited
95 geographical scope, which makes it difficult to draw comprehensive conclusions across
96 different climate zones. To better understand the epidemiology and ecology of seasonal
97 variation in mortality and the role of temperature, a systematic and comprehensive
98 investigation on highly diverse populations including multiple locations from multiple climate
99 zones is crucial. Such investigation should further help us to develop hypothesis about the
100 impact of warming climate on seasonal dynamics of mortality for future investigations.

101 Moreover, the magnitude of seasonal variation in mortality varies substantially among different
102 locations, possibly due to the differences in location-specific characteristics, e.g., socio-
103 economic development. To date, some studies have explored this issue but were limited in
104 geographical locations and climate zones.^{5,9,10,13–16} A comprehensive evaluation across
105 multiple locations with various characteristics is warranted, as it will aid in identifying more
106 vulnerable locations with a greater need for intervention.

107 In this research, we investigated seasonality of mortality, with particular focus on its
108 magnitude, by analyzing daily time-series data of mortality and temperature from 719 locations

109 (i.e., city/province/prefecture) in 34 countries from tropical, dry, temperate and continental
110 climate zones. Our primary focus in this study was to estimate the magnitude of seasonal
111 variation in mortality (i.e., seasonal amplitude) and to examine the extent to which temperature
112 explains seasonality of mortality. We also evaluated the modifying effects of location-specific
113 characteristics on seasonal variation of mortality. To our knowledge, this is by far the largest
114 multi-country study on seasonality of mortality.

115 **Methods**

116 Data collection

117 We collected daily time-series of mortality and mean temperature from 719 locations in 34
118 countries in largely overlapping periods ranging from January 1969 to December 2016. The
119 data were obtained through the Multi-country Multi-city (MCC) Collaborative Research
120 Network (<http://mccstudy.lshtm.ac.uk/>).¹⁷ Mortality was represented by daily counts of death
121 from all causes or, where not available, non-external causes (International Classification of
122 Diseases [ICD]-9 0-799, ICD-10 A00-R99) and cardiovascular (ICD-8 390-458, ICD-9 390-
123 459, ICD-10 I00-I99) and respiratory diseases (ICD-8 and ICD-9 460-519, ICD-10 J00-J99).
124 We also obtained information on Köppen–Geiger climate groups for each location, including
125 tropical, dry, temperate, and continental climate zones.¹⁸

126 We collected data on the indicators of location-specific characteristics for each location,
127 including environmental factors, demographics, and socioeconomic factors. For environmental
128 indicators, we considered the multi-year average value of daily mean temperature, daily mean
129 temperature range, daily mean relative humidity, and annual PM_{2.5} levels. For demographic
130 and socioeconomic factors, we collected from the Organisation for Economic Co-operation and
131 Development Regional and Metropolitan Database, including the proportion of population
132 aged over 65 years old, gross domestic product (GDP), gross value added (GVA, a measure of

133 labour productivity), education level, unemployment rate, and Gini index (a measure of wealth
134 inequality). The details for the indicators are included in supplementary material (Page 5,
135 supplementary material).

136 Statistical analysis

137 *Estimating location-specific seasonality*

138 In the first step, we performed location-specific time-series analyses to assess seasonality of
139 mortality using quasi-Poisson regression models¹⁹ throughout the study period available in
140 each location. Day-of-year was considered as the exposure indicator for seasonality. This is
141 different from previous studies,²⁻¹⁰ which used monthly aggregated data to compare winter
142 mortality with other times of the year. In this study, we took values from 1 to 366 to represent
143 day of year, corresponding to 1 January through 31 December for locations in the northern
144 hemisphere and 1 July to 30 June of the following year for locations in the southern hemisphere.
145 To model seasonality we used a cyclic spline with 4 degrees of freedom (*df*) for day of year.
146 The days-of-year with maximum and minimum mortality predictions were identified as the
147 peak and trough, respectively, of seasonality of mortality. We then took the ratio of mortality
148 predicted at peak to mortality prediction at trough (peak-to-trough ratio, PTR) to summarise
149 seasonality. A stratum defined by year, day of week and their interaction was used to control
150 for long-term trends and effect of day of week.

151 We then added temperature to the model described above for each location, by using a
152 distributed lag non-linear model (DLNM)²⁰ to estimate seasonality adjusting for temperature
153 effect. We modelled the non-linear and non-linearly delayed effect of temperature on mortality
154 using a cross-basis with natural cubic spline for temperature with three internal knots at the
155 25th, 50th, and 75th percentiles of temperature, and another natural cubic spline for lag with 3

156 *df*. The lag was extended up to 21 days.¹⁷ From this model, we also calculated PTR to represent
157 temperature-adjusted seasonality.

158 *Pooling the location-specific seasonality by country and climate zone*

159 In the second step, location-specific estimates of seasonal curve (i.e., coefficients of knot
160 points) with and without temperature adjustment were pooled separately by climate zone
161 through two-level (locations nested within country/region) random-effects multivariate meta-
162 analysis techniques.²¹ We also pooled the estimates by each of 34 countries/regions with
163 location considered as the random effect factor. Using the pooled coefficients, the seasonal
164 curve and corresponding PTR were estimated for each country and each climate zone.

165 *Modification of seasonality by location-specific characteristics*

166 In the final step, we first explored the between-location heterogeneity of the seasonal curve by
167 including location-specific average temperature, temperature range, indicator for country, and
168 indicator for Köppen–Geiger climate zone as meta-predictors in the random effects meta-
169 regression.²² The heterogeneity was tested for location-specific seasonality estimates before
170 and after adjusting for temperature separately. Next, we evaluated the association of unadjusted
171 and adjusted PTR with each indicator in separate meta-regression models including indicators
172 for countries and climate zones. For each indicator, the original value was scaled by the
173 country’s average value to remove the between-countries effects from the correlation. Results
174 were expressed as log (PTR) variation for a standard deviation increase of the indicator.

175 *Sensitivity Analysis*

176 We performed several sensitivity analyses. First, we evaluated how results changed with 5 and
177 6 *df* for the cyclic spline included in the location-specific time-series regression model.
178 Second, we conducted seasonality assessment by using the subset of data since year 2000.

179 Finally, we investigated the sensitivity by changing the types of splines, lag days and df for the
180 cross-basis function for temperature adjustment in the location-specific regression model.

181 We investigated seasonality of all-cause, cardiovascular, and respiratory mortality in separate
182 analyses with R software, version 3.6.0 (R Development Core Team) using `dlm` and `mixmeta`
183 packages.

184 **Results**

185 The final analysis included 138 868 448 deaths from all or non-external causes in 719 locations
186 in 34 countries, 39 777 149 deaths from cardiovascular diseases, and 12 805 050 deaths from
187 respiratory mortality in 519 locations in 22 countries. The country-specific average mean
188 temperature ranged from 4.7°C in Norway to 27.6°C in Thailand. These temperatures are
189 illustrative of locations characterised by four Köppen–Geiger climate zones¹⁸ (Figure 1),
190 including 94 locations in the tropical climate zone (e.g., Ho Chi Minh City, Vietnam), 57
191 locations in the dry climate zone (e.g., Mashhad, Iran), 440 locations in the temperate climate
192 zone (e.g., London, UK), and 128 locations in the continental climate zone (e.g., Hokkaido,
193 Japan). Table 1 shows a summary of daily data for each climate zone. Supplementary Table S1
194 summarises mortality in each season for each country/region.

195 A descriptive summary of location-specific indicators is shown in Supplementary Table S2.
196 Unemployment rate and PM_{2.5} concentrations showed a large variation between locations.
197 Socioeconomic indicators included in the analysis were correlated, and averaged mean
198 temperature was correlated with the other indicators (Supplementary Figure S1).

199 Before adjustment for temperature, a seasonal pattern was observed in all climate zones with a
200 high mortality in cold seasons and a low mortality in warm seasons (Figure 2). When
201 temperature was adjusted, seasonality of mortality remained higher in cold seasons in most
202 climate zones, except for seasonality of all-cause mortality in the tropical climate zone, where

203 the adjusted seasonality became almost flat with a large confidence interval (Figure 2). The
204 unadjusted PTR varied between climate zones, with the lowest estimate observed in the tropical
205 zone (Table 2). The unadjusted PTRs for all-cause mortality were 1.05 (95% confidence
206 interval (CI) 1.00–1.11) in the tropical climate zone and 1.23 (95% CI: 1.20–1.25) in temperate
207 climate zone, respectively (Table 2). Adjusting for temperature reduced the PTRs to different
208 degrees, from a slight reduction observed in the tropical climate zone to a large reduction in
209 the temperate climate zone: the pooled unadjusted PTR for all-cause mortality was reduced to
210 1.02 (95% CI: 0.95–1.09) in tropical climate zone, and 1.10 (95% CI: 1.07–1.12) in temperate
211 climate, respectively (Table 2).

212 Our findings were generally similar for cause-specific mortalities (Figure 2 and Table 2).
213 However, the change of seasonality estimates by temperature adjustment was more evident for
214 cardiovascular mortality while less profound for respiratory mortality.

215 The location-specific seasonality estimates were presented in Figure 3 (Supplementary Table
216 S3) and summarised for each country/region in Figure 4 (Supplementary Figure S2). Although
217 the location- and country/region-specific results are generally consistent with our findings from
218 the climate zone-specific assessment, PTR estimates varied between
219 locations/countries/regions even for those within the same climate zone (Page 37,
220 supplementary material).

221 Our meta-analysis showed substantial heterogeneity between locations for seasonality
222 estimates both with and without temperature adjustment (Supplementary Table S4). Results
223 from our multivariate meta-regression models suggest that the heterogeneity for seasonality
224 estimates with/without temperature adjustment for all mortality types was reduced when
225 country indicators were included (Supplementary Table S4). The other three predictors (i.e.,
226 the indicators for climate zones, location-specific average mean temperature and location-
227 specific total temperature range) all significantly modify the effect of seasonality both in the

228 single-predictor and the full models, and account for a small proportion of heterogeneity
229 (Supplementary Table S4).

230 Figure 5 presents the association between each of location-specific characteristics and PTR.
231 Average mean temperature was positively associated with unadjusted PTR for all-cause
232 mortality, and adjusting for temperature in PTR moved the estimate towards the null. The other
233 indicators showed no associations with PTRs for all-cause mortality. Our analysis of cause-
234 specific mortality showed similar results, with a few additional findings: for cardiovascular
235 mortality, total range of daily mean temperature was negatively associated with unadjusted
236 PTR, which moved toward the null after adjustment for temperature in PTR; for respiratory
237 mortality, averaged mean relative humidity was negatively associated with both unadjusted
238 and adjusted PTR.

239 Results from sensitivity analyses (Supplementary Table S5) suggest that pooled seasonality
240 curve and PTR in each climate zone of the main analysis were generally robust to different
241 approaches (Supplementary Figure S3 & Figure S4). Country- and region-specific PTR
242 estimates for those with most locations characterised by tropical climates seemed to be less
243 sensitive to different modelling choices (Supplementary Table S6 and Table S7). Unadjusted
244 PTR for most countries/regions was reduced in the subperiod analysis by using data since year
245 2000 (Supplementary Table S6). In addition, associations between indicators and PTR
246 remained similar when using different approaches (Supplementary Figure S5).

247 **Discussion**

248 Our study systematically and comparatively investigated seasonality of mortality in 719
249 locations of 34 countries covering a wide range of environmental conditions, population
250 dynamics and socioeconomic status. To our knowledge, this is the largest investigation on
251 seasonality of mortality. Our study provides evidence that the generally higher mortality in

252 cold seasons than in warm seasons is considerably explained by temperature, and this pattern
253 is most evident in temperate and continental climate zones. Despite a similar pattern, the
254 amplitudes of seasonality varied between locations. Locations characterised by warm climate
255 experienced larger seasonal variations in mortality, which was related to stronger effect of
256 temperature. Our investigation of this long-known complex phenomenon provides important
257 evidence for understanding this phenomenon and informing the ongoing discussion on future
258 impacts of warming climate.

259 Winter peaks and summer troughs in seasonality of mortality have been broadly defined and
260 consistently described in previous studies,²⁻¹⁰ and we observed a similar seasonal pattern for
261 most of the locations in our study. Although previous studies measured the magnitude of
262 seasonality in mortality, direct comparison with our findings (i.e., unadjusted PTR) is difficult
263 due to the differences in modelling approaches. Where we applied time-series analysis to
264 estimate mortality on each day of the year and then compared maximum mortality estimates
265 with minimum mortality estimates on a daily basis to measure the strength of seasonality,
266 previous studies used mortality data aggregated to each month or, to a lesser extent, for each
267 week, and applied Fourier transforms to compare mortality estimates in peak months with those
268 trough months. Stewart et al. reviewed 48 studies on seasonality of cardiovascular mortality
269 mostly from temperate areas in Europe and North America and reported an estimate of 1.23-
270 fold (95% CI: 1.16–1.31) for the relative difference of cardiovascular mortality in peak-versus-
271 trough season,² which was lower than our estimate on seasonality of cardiovascular mortality
272 in temperate zone (1.32 (95% CI: 1.27–1.36)).

273 One highlight of our investigation is the assessment of the extent to which the seasonal
274 variation of temperature is associated with seasonal variation in mortality. Despite the
275 extensive literature on the effects of cold and hot temperatures on health, debate remains
276 regarding whether temperature is the main cause for seasonality of mortality.^{11,12,23} Addressing

277 this issue is essential for understanding the epidemiology and ecology of seasonal variation in
278 mortality. Using multi-decade data from 36 cities in the US and three cities in France covering
279 a wide range of winter temperatures from -5 to over 20°C , Kinney et al. observed no
280 correlations between seasonal temperature differences (the difference in mean temperature
281 between winter and summer) and winter excess mortality, and concluded that temperature was
282 not a key driver of winter excess mortality.¹² However, this conclusion can be misleading, as
283 their findings actually answered the question whether the spatial variation in the strengths of
284 seasonal variation in mortality was related to the differences in seasonal temperature
285 differences. In our study, we estimated temperature-adjusted seasonal variation in mortality
286 and demonstrated that temperature is an important driver of seasonal variation in mortality,
287 especially in temperate/continental climate zones. Our findings, on the other hand, provide a
288 basis for developing hypotheses about the potential impact of climate change on seasonality of
289 mortality, for example, whether an increasing temperature and shortening winter season will
290 reduce winter mortality, increase summer mortality, and subsequently attenuate their variation
291 between seasons. Future investigations are merited to investigate these hypotheses by taking
292 into account the increasing extreme weathers (e.g., cold spells, snowfall or ice), other seasonal
293 events (e.g., infectious disease outbreaks) and human adaptation, which is beyond the scope of
294 the current study.

295 It should be noted that other unmeasured seasonally varying factors, e.g., sunlight, rainfalls,
296 infectious disease incidence, and human behaviour, may also contribute to seasonality of
297 mortality.² For example, the increase in infectious disease-related mortality during rainy season
298 may explain seasonal variation in total mortality in the tropical climate zone,¹³ and influenza
299 infections may increase the risk of excess mortality in winter.²³⁻²⁵ Furthermore, we found that
300 seasonal variation in respiratory mortality seems to be less explained by temperature than are
301 all-cause and cardiovascular mortality. This result may be explained by the fact that the

302 increase in respiratory mortality during the winter season can be considerably attributed to
303 seasonal respiratory infections (e.g., influenza and respiratory syncytial virus). Further research
304 in seasonal pattern of mortality considering various kinds of seasonally varying factors would
305 complement the evidence provided in this study.

306 Our results showed a significant spatial variation in the amplitude of seasonality across
307 locations, and climate factors at location level contributed to this spatial variation but cannot
308 fully characterise differences between locations. Before adjusting for temperature in
309 seasonality assessment, we found a larger seasonal variation in locations characterised by warm
310 climate; this modification became weak on the remaining seasonality after removing the short-
311 term effect of temperature. Consistently, previous studies on the effects of cold temperature
312 reported that cold-related mortality was higher in warm climates than in cold climates.^{10,26,27}

313 One explanation is that populations routinely exposed to warm climate are less adapted to or
314 prepared for cold weather during the year (e.g., lack of proper insulation). In addition, our
315 results in cause-specific mortality showed that populations from less humid areas may exhibit
316 a large seasonal variation for respiratory mortality. This result may be related to the impact of
317 humidity on respiratory tract infections and transmissions (e.g., fomites). Low humidity in cold
318 weather may increase survival of influenza virus and increase its transmission,²⁴ and a decrease
319 in temperature and humidity can precede the onset of infections.²⁸ Therefore, humidity can
320 possibly modify seasonality of respiratory mortality. Elaborating on this phenomenon could be
321 a topic for future studies.

322 Some limitations must be acknowledged. First, our seasonality assessment was based on the
323 assumption that seasonal variation in mortality and the role of temperature have not changed
324 over the study period. In our sensitivity analysis, we repeated the assessment by using the data
325 since 2000: although the results showed a reduction in unadjusted PTR for most
326 countries/regions, the main findings and conclusions did not change. However, future studies

327 are warranted to investigate this complex research topic— whether or not and how seasonality
328 of mortality has changed over the years. Second, we used PTR as a numeric measure of
329 seasonality, which may be limited as it only quantifies the amplitude of seasonal variation in
330 mortality. In other words, PTR is not able to reflect the shape of seasonal variation in mortality.
331 Further investigations would be beneficial by improving seasonality assessment, e.g.,
332 quantifying the area under seasonal curve as attributable fraction. Third, coverage of tropical
333 and dry climate zones and less developed locations was limited in our study, especially for
334 cardiovascular and respiratory mortality, so the results for these areas should be interpreted
335 with cautions. The country-level estimates for several countries (e.g., Sweden, China and Iran)
336 may not be representative, as only a small number of locations from these countries were
337 included in our analysis. Fourth, we did not explore modifying effect of indicators by using a
338 multivariable model, because of a high correlation between indicators. Finally, the collection
339 (e.g., case ascertainment, codification) and processing of mortality data may vary between
340 countries.

341 Despite these limitations, our study is, to our knowledge, the largest investigation on
342 seasonality of mortality. This multi-country study used the largest database of location-level
343 daily time-series for mortality for 719 locations from 34 countries and identified a strong
344 seasonal variation in mortality in temperate climate zones, which was attenuated substantially
345 after adjusting for temperature, whereas a small seasonal variation was observed in tropical
346 climate zone. Moreover, populations consistently exposed to warm climates seem to be more
347 susceptible to seasonal variation in mortality. Based on this large and geographically versatile
348 dataset and well-tested methods, our findings provide a better understanding of this long-
349 known complex phenomenon and a basis for generating hypotheses about the future impact of
350 climate change on seasonality of mortality, which ultimately could help with the development
351 of health systems and infrastructure planning in the future.

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Ethics approval: Not required.

Data availability:

Data have been collected within the MCC (Multi-City Multi-Country) Collaborative Research Network (<https://mccstudy.lshtm.ac.uk>) under a data sharing agreement and cannot be made publicly available. The R code for the analysis is available from the first author.

Supplementary data:

Supplementary data are available at IJE online.

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Conflict of interest:

None declared.

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Figure legends:

Figure 1. Spatial distribution of location-specific averaged annual mean temperature of 719 locations in four Köppen–Geiger climate zones (A: Tropical, B: Dry, C: Temperate, and D: Continental)

Figure 2. Seasonality of mortality without (black) and with (red) temperature adjustment in four Köppen–Geiger climate zones (A: Tropical, B: Dry, C: Temperate, and D: Continental)

The seasonality is computed as the relative risk (RR) of mortality estimates at each day-of-year to daily minimum mortality estimates at the trough day with 95% confidence intervals (95% CIs) for four Köppen–Geiger climate zones:

$$\text{Relative risk} = \frac{\text{Mortality estimate at day}_i}{\text{Minimum mortality estimate at the trough}}$$

These estimates are obtained by pooling location-specific estimates for each climate zone. We took values from 1 to 366 to represent day of year, corresponding to January 1st through December 31st for locations in the northern hemisphere and July 1st to June 30th of the following year for locations in the southern hemisphere (for common years, values were taken from 61 to 366 from the 60th day to the 365th day).

Figure 3. Peak-to-trough ratio (PTR) with 95% confidence intervals (95%CI) without (left) and with (right) temperature adjustment for each location for all-cause/non-external (blue), cardiovascular (red), and respiratory (green) mortality

The size of the points corresponds to the precision of the PTR estimate (i.e., the inverse of the standard error of the PTR).

Figure 4. Peak-to-trough ratio (PTR) with 95% confidence intervals (95%CI) without (black) and with (red) temperature adjustment for each country/region (numbers of locations in each country/region for each Köppen Geiger climate zone[§])

These estimates are obtained by pooling location-specific estimates for each country/region.

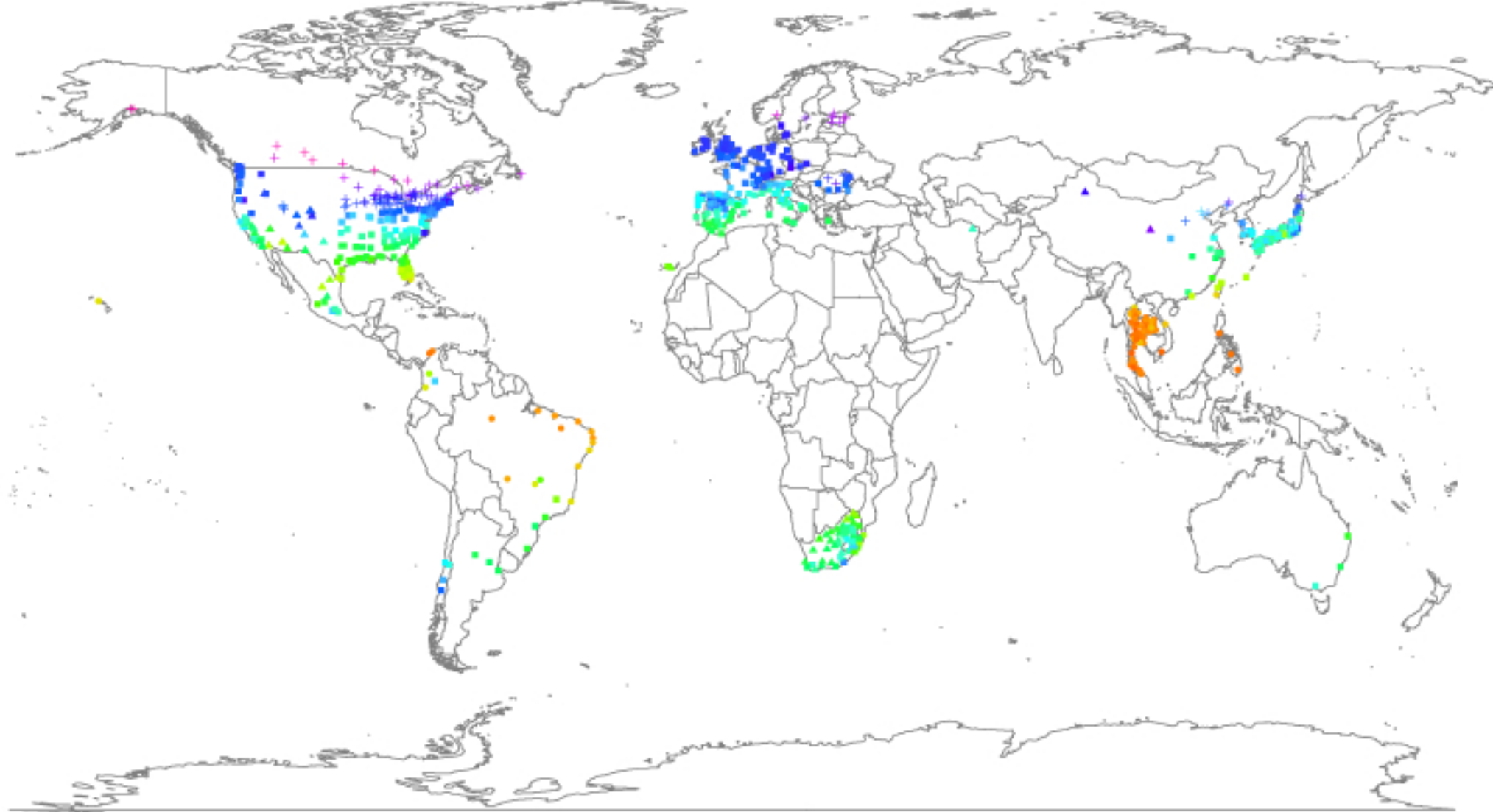
**Countries/regions which have data for all mortality causes.*

§ Four Köppen Geiger climate zones (A: Tropical, B: Dry, C: Temperate, and D: Continental)

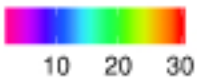
Different background colors were used to highlight the climate zone for each country (red: Tropical, yellow: Dry, green: Temperate, blue: Continental, and grey: multiple climate zones).

Figure 5. Associations between the indicators on location-specific characteristics and peak-to-trough ratio before (black) and after (red) temperature adjustment.

Coefficients with 95% confidence intervals (95% CIs) were obtained from a meta-regression model adjusted by indicators for country and climate zone. Results are expressed as the changes in $\log(\text{PTR})$ for standard deviation increase in the indicators.



Averaged annual
mean temperature(°C)



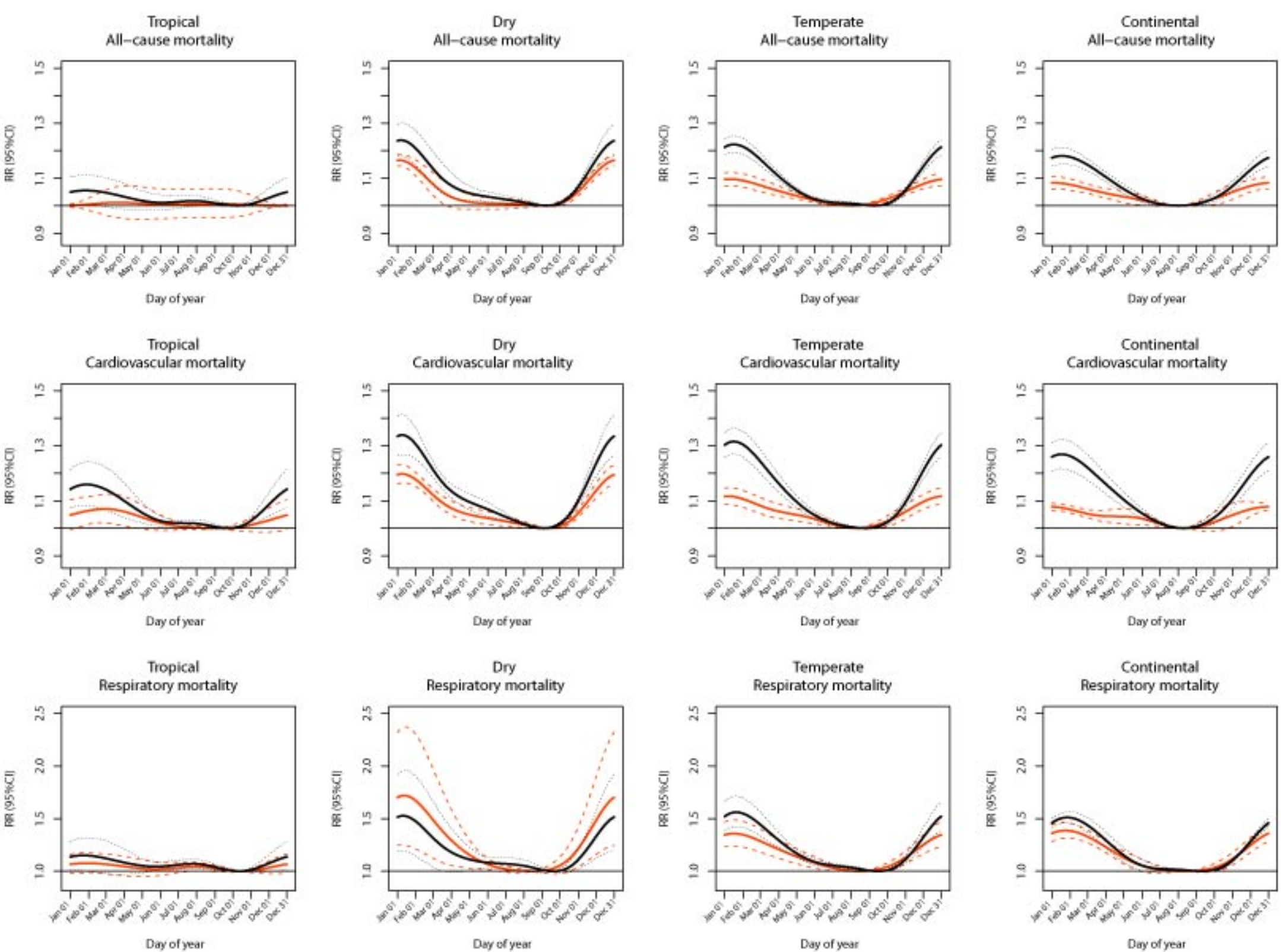
Köppen-Geiger climate

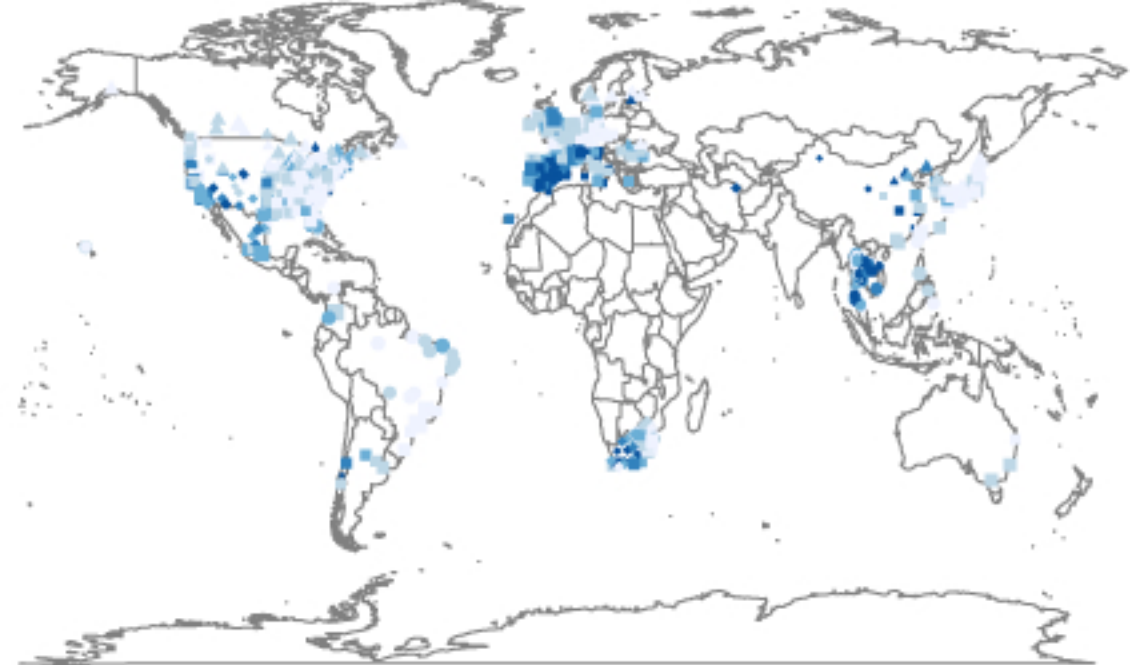
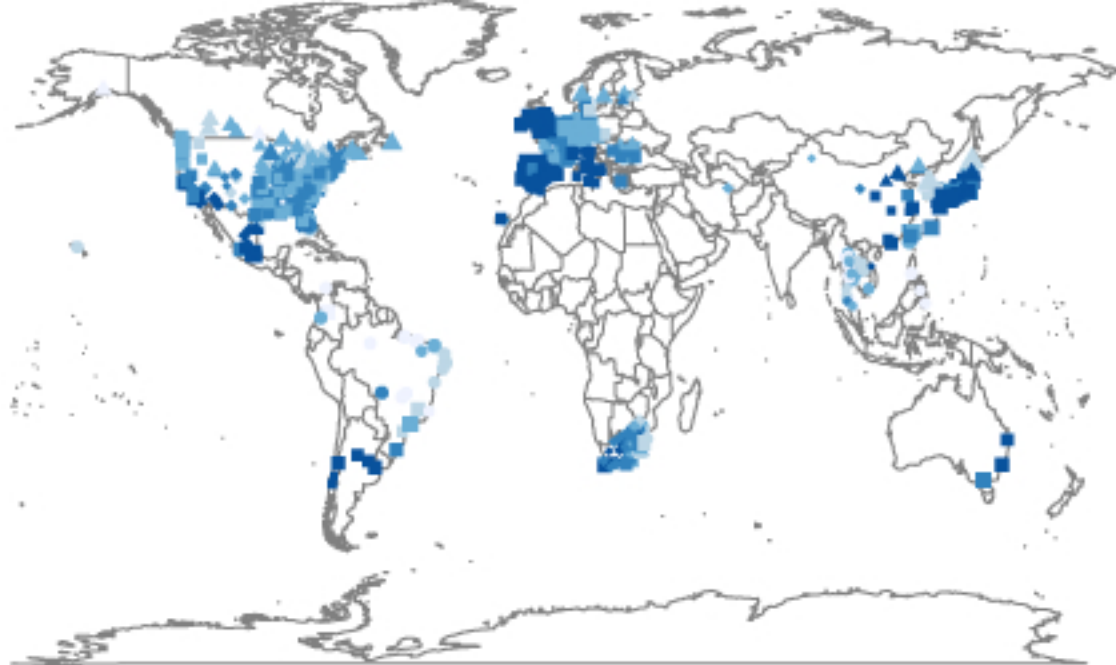
• A(tropical)

▲ B(dry)

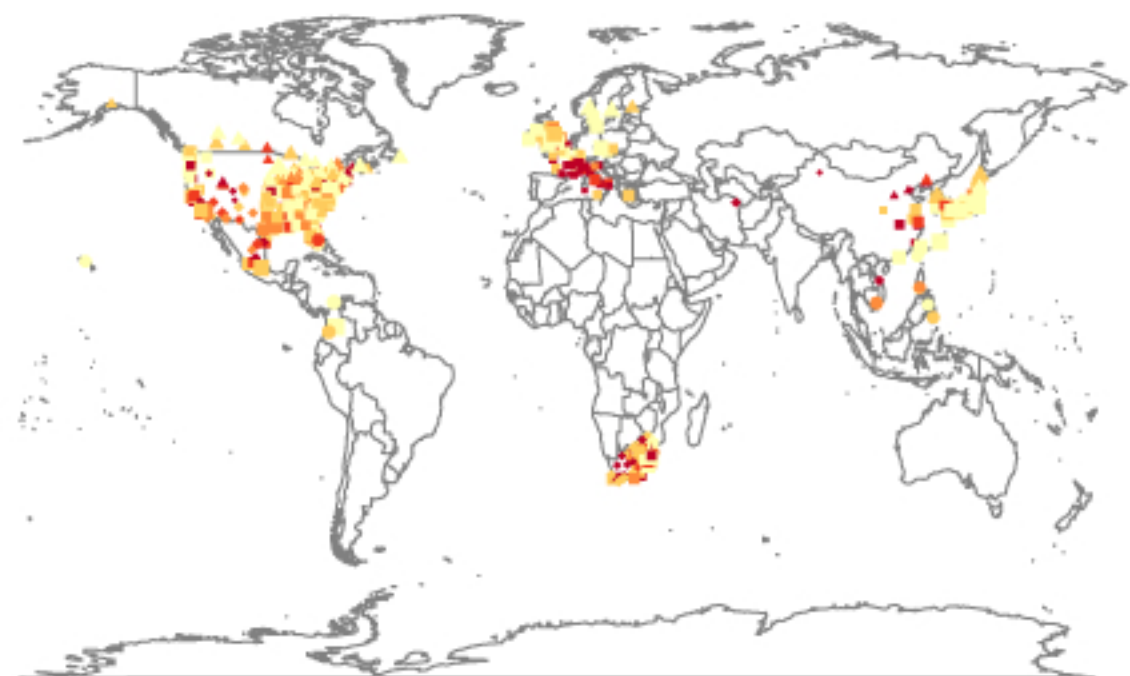
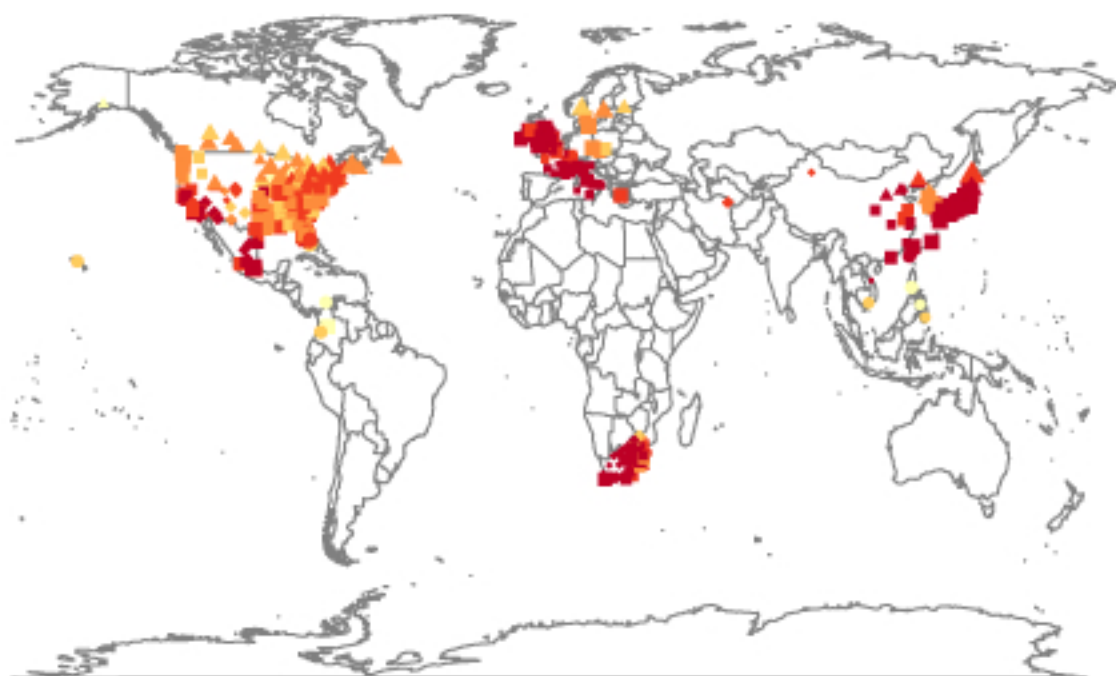
■ C(Temperate)

+ D(continental)

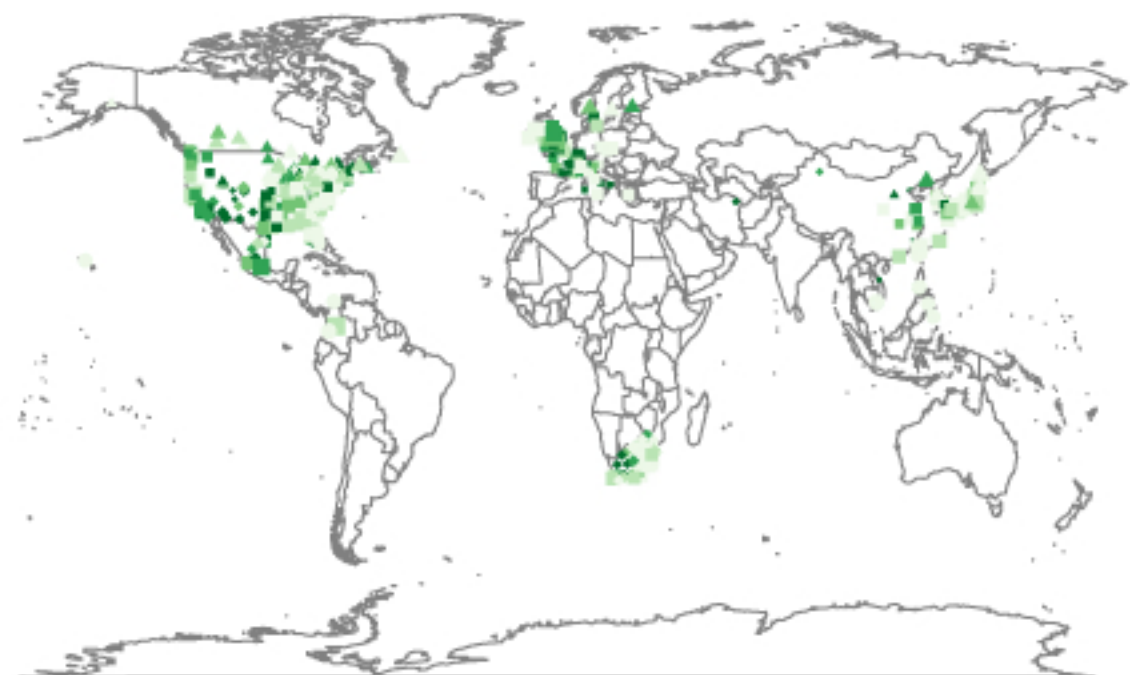
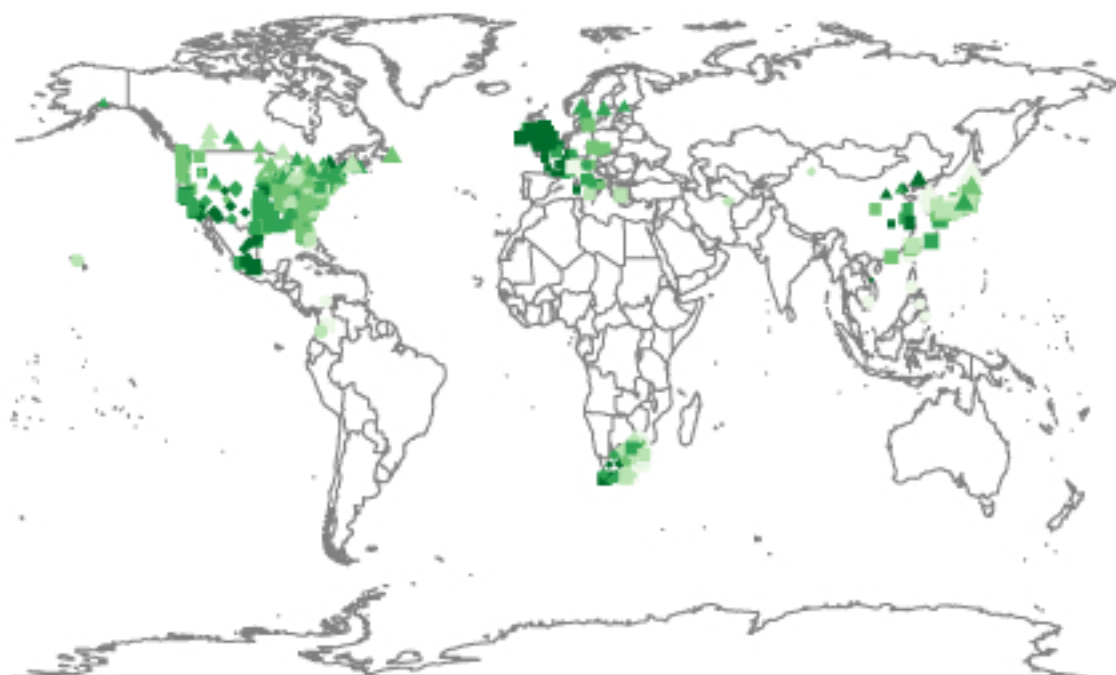




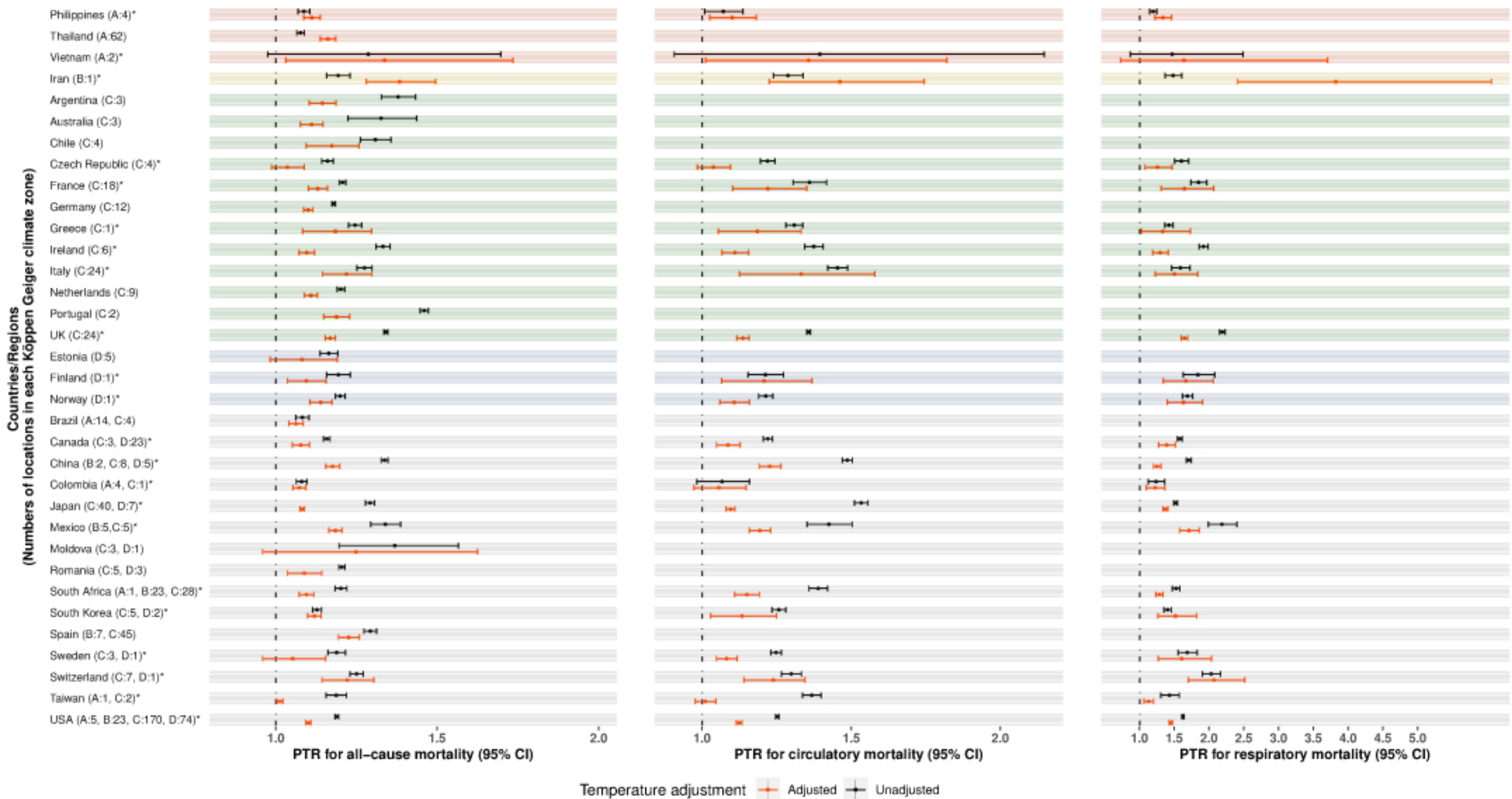
Köppen-Geiger climate ● A(tropical) ◆ B(dry) ■ C(temperature) ▲ D(continental) Peak-to-Trough Ratio (PTR) <1.11 1.15 1.20 1.26 >1.26 Standard Error ● 0.05 ● 0.02 ● 0.01



Köppen-Geiger climate ● A(tropical) ◆ B(dry) ■ C(temperature) ▲ D(continental) Peak-to-Trough Ratio (PTR) <1.13 1.21 1.26 1.34 >1.34 Standard Error ● 0.05 ● 0.02 ● 0.01



Köppen-Geiger climate ● A(tropical) ◆ B(dry) ■ C(temperature) ▲ D(continental) Peak-to-Trough Ratio (PTR) <1.39 1.54 1.65 1.84 >1.84 Standard Error ● 0.20 ● 0.10 ● 0.05



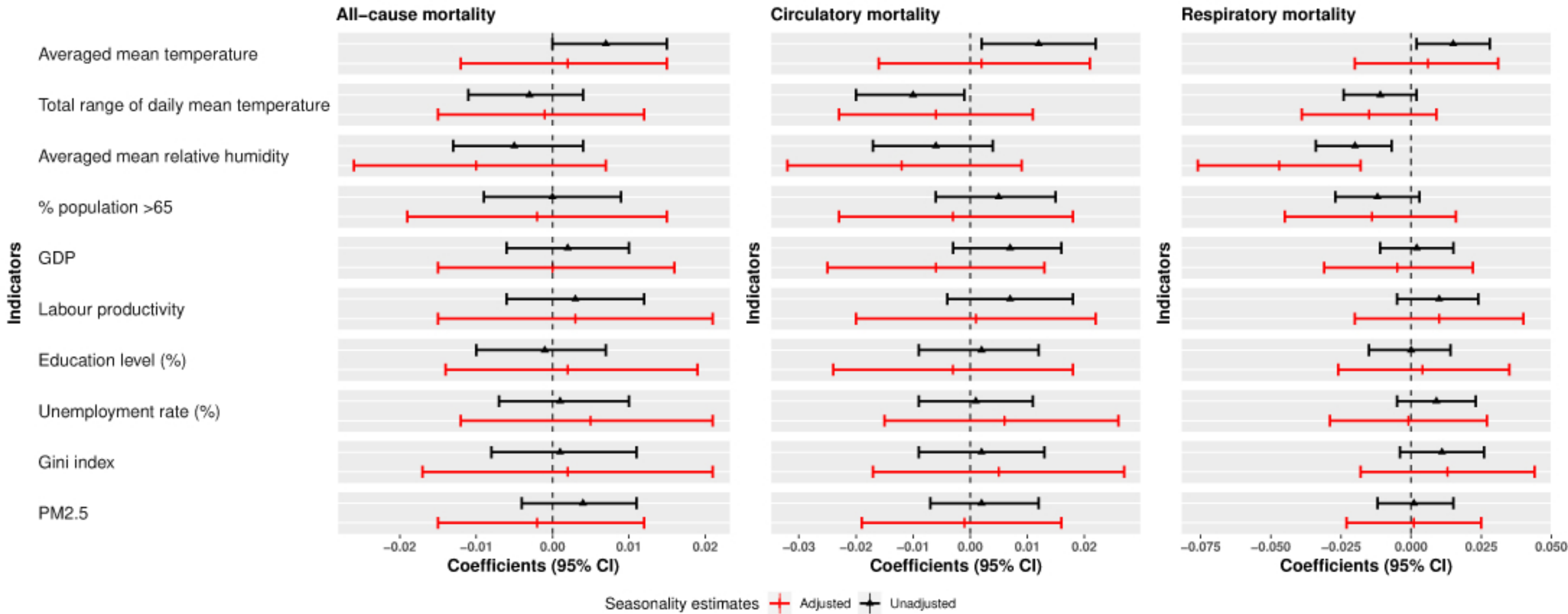


Table 1. Summary (mean± standard deviation) of daily mean temperature (°C) and daily mortalities (counts) by climate zones

Climate Zone	No. of locations [†]	Mean temperature	All-cause mortality [*]	Cardiovascular mortality	Respiratory mortality
Tropical	94/18	26.68±2.86	14.91±14.31	8.82±6.05	2.44±2.26
Dry	57/50	17.61±8.17	13.82±12.87	3.81±4.55	1.68±2.19
Temperate	440/350	14.51±8.43	23.82±36.96	8.44±15.19	2.58±5.34
Continental	128/118	8.85±10.87	11.58±17.86	4.24±6.41	0.96±1.37

* Data on non-external mortality was used when data on all-cause mortality is not available for some locations.

[†]No. of locations where all-cause/non-external mortality data are available/ No. of locations where cause-specific mortality data are available

Table 2. Pooled peak-to-rough ratio (95% confidence intervals) for each climate zone.

	Temperature	Tropical	Dry	Temperate	Continental
All-cause mortality	Unadjusted	1.05 (1.00, 1.11)	1.23 (1.18, 1.30)	1.23 (1.20, 1.25)	1.20 (1.17, 1.23)
	Adjusted*	1.02 (0.95, 1.09)	1.16 (1.14, 1.19)	1.10 (1.07, 1.12)	1.08 (1.06, 1.10)
Cardiovascular mortality	Unadjusted	1.16 (1.08,1.24)	1.34 (1.27,1.41)	1.32 (1.27,1.36)	1.27 (1.22,1.32)
	Adjusted*	1.07 (1.01, 1.13)	1.20 (1.16, 1.23)	1.11 (1.10, 1.13)	1.08 (1.07, 1.10)
Respiratory mortality	Unadjusted	1.19 (1.07, 1.33)	1.53 (1.19, 1.95)	1.61 (1.42, 1.73)	1.55 (1.46, 1.66)
	Adjusted*	1.08 (0.99, 1.17)	1.72 (1.25, 2.37)	1.36 (1.24, 1.49)	1.39 (1.31, 1.46)

* Temperature was adjusted for each location by using a distributed lag non-linear model (DLNM): the non-linear exposure-response association was modelled by a natural cubic spline function with three internal knots at 25th, 50th, and 75th percentiles of temperature, and the lag-response curve was fit by another natural cubic spline function with 3 df with extended lag up to 21 days.