

Upjohn Institute Working Papers

Upjohn Research home page

12-2-2022

Salary History and Employer Demand: Evidence from a Two-Sided Audit

Amanda Agan Rutgers University, amanda.agan@rutgers.edu

Bo Cowgill Columbia University, bo.cowgill@gsb.columbia.edu

Laura K. Gee *Tufts University*, laura.gee@tufts.edu

Upjohn Institute working paper ; 22-379

Follow this and additional works at: https://research.upjohn.org/up_workingpapers

Part of the Labor Economics Commons

Citation

Agan, Amanda, Bo Cowgill, and Laura K. Gee. 2022. "Salary History and Employer Demand: Evidence from a Two-Sided Audit." Upjohn Institute Working Paper 22-379. Kalamazoo, MI: W.E. Upjohn Institute for Employment Research. https://doi.org/10.17848/wp22-379

This title is brought to you by the Upjohn Institute. For more information, please contact repository@upjohn.org.

Salary History and Employer Demand: Evidence from a Two-Sided Audit

Upjohn Institute Working Paper 22-379

Amanda Agan Rutgers University amanda.agan@rutgers.edu Bo Cowgill Columbia University bo.cowgill@gsb.columbia.edu Laura Gee Tufts University laura.gee@tufts.edu

December 2022

ABSTRACT

We study how salary history disclosures affect employer demand, and how salary history bans shape hiring and wages. We show how these effects depend on the properties of the labor market, and we measure the key properties using a novel, two-sided field experiment. Our field experiment features hundreds of recruiters reviewing more than 2, 000 job applications. We randomize the presence of salary history questions as well as job candidates' disclosures. We find that employers make negative inferences about nondisclosing candidates, and that they anticipate positive selection into disclosure. Recruiters view salary history as a stronger signal about competing options than about worker quality. Disclosures by men (and other highly paid candidates) yield higher salary offers; however, they are negative signals of the value (net of salary) that a worker brings to the firm, and thus they yield fewer callbacks. While disclosures (especially of high amounts) generally increase recruiter beliefs about quality and competing offers, male wage premiums are regarded as a weaker signal of quality than other sources (such as the premiums from working at higher-paying firms or being well paid compared to peers). Recruiters correctly anticipate that women are less likely to disclose salary history at any level, and thus they punish women less than men for silence. When we simulate the effect of salary history bans using our results, we find muted effects on callbacks. Gender inequality in salary offers is reduced; however, equality comes at the expense of lower salaries overall (and especially for men).

JEL Classification Codes: J71, M51, C93

Key Words: blinding, disclosure, recruiting, statistical discrimination, salary history, field experiments

Acknowledgments: The authors thank David Abrams, Kyle Coombs, Gordon Dahl, Katie Donovan, Peter Kuhn, Joanna Lahey, Nina Roussille, Bobby Pakzad-Hurson, Patryk Perkowski, Lamar Pierce, David Reiley, Jonah Rockoff, Chris Stanton, Olga Stoddard, Olga Shurchov, and Laurina Zhang for valuable feedback and comments, as well as participants at the MEER Conference, the Advances in Field Experiments Conference, the American Economic Association Annual Meeting, the Innovation and Institutions Conference, the Economic Science Association Conference, the Empirical Management Conference, Entrepreneurship and Private Enterprise Development (EPED), NBER Market Design, the NBER Summer Institution (Labor), the SMS Annual Conference, the Wharton People Analytics Conference, and the seminars at Barnard College, Boston College, Clemson University, Columbia University, the Federal Reserve Board, Florida State University, George Mason University, Nuffield College (Oxford), Simon Fraser University, SUNY Albany, the University of California, Santa Barbara, the University of California Santa Cruz, the University of Connecticut, the University of Missouri, the University of Pennsylvania Law School, the University of Texas at Dallas, and the University of Illinois at Urbana-Champaign (UIUC). We also thank Hailey Brace, Karishma Chouhan, Nadine Fares, Zachary Finn, Matt Fondy, and Norman Yuan for research assistance. And we acknowledge financial support from the W.E. Upjohn Institute and Facebook Research. The experiments in this paper were approved by an Institutional Review Board and preregistered at the AEA RCT Registry under AEARCTR-0003088.

Upjohn Institute working papers are meant to stimulate discussion and criticism among the policy research community. Content and opinions are the sole responsibility of the author.

1 Introduction

Since 2016, 21 states and 21 jurisdictions in the United States have banned employers from asking job applicants about their salary histories.¹ The goal of these laws is to ameliorate historical inequalities relating to gender, race/ethnicity, or other protected characteristics. However, the information encoded in historical salaries is complex. Salary history is provided voluntarily, and employers may anticipate strategic aspects of disclosure. When revealed, salary histories may signal subtle information about candidates' hidden characteristics and outside options. Disclosure behaviors may differ between groups of candidates, requiring complex inferences by employers.

How employers use salary-history signals—and what inferences they make in their absence—is the subject of this paper. We study how salary history disclosures affect employer demand and how salary history bans shape hiring and wages. Our conceptual model shows how these effects depend on the characteristics of the labor market. We then measure the key underlying properties—and directly measure the effects of disclosures and bans—in a novel, large-scale field experiment.

Our analysis is focused on three central questions: First, what inferences do employers make about candidates who refuse to disclose salary history? Second, among candidates who do disclose it, what do employers infer from higher or lower salary amounts? Finally, how do employers combine these inferences into decisions about whom to hire and how much to pay them? Our conceptual framework shows how these answers depend on two key properties: 1) how correlated salary history amounts are with workers' quality and outside offers, and 2) whether a high previous salary is correlated with low value (net cost) to the employer, or vice versa.

To answer these questions empirically, we developed a novel, two-sided field experiment. Using an intermediary firm, we assumed the role of an employer and hired hundreds of recruiters to make decisions on more than 2,000 job applications for a software engineering position. The recruiters in our field experiment were paid real wages and faced real incentives, but they were not aware that our job applicants were fictitious. On the demand side, we randomly vary the presence of a salary-history prompt on the job application form; and, on the candidate side, we randomly vary whether candidates disclose their current salaries. Among candidates who disclose them, we also vary the levels of salaries, mimicking real-world, male-favoring gender wage gaps.

We then measure recruiters' callback recommendations and salary offers for all candidates. By examining different bargaining protocols, we also infer recruiters' maximum salary offer (e.g., willingness-to-pay, or WTP) for all candidates, as well as their beliefs about each worker's competing offers. To contextualize these findings, we present the results of a large survey of the U.S. workforce that measured the prevalence of salary history questions and candidates' responses to

¹See https://www.hrdive.com/news/salary-history-ban-states-list/516662/ for the most up-to-date list.

them.

Our field experiment yields three main results. First, employers make negative inferences about "silent" candidates. Candidates who do not disclose their salary are assumed to have lower-thanaverage quality and lower outside options, and are given lower salary offers. Disclosure choices have a large impact on recruiters' beliefs about candidates' outside offers, and a smaller impact on recruiters' beliefs about candidates' quality (measured by WTP)—but both are affected in the same direction. We find that women are punished less for silence. Our survey suggests a possible reason: women at all salary levels are less comfortable disclosing their salary, which makes their choice to disclose less informative about value.

Our second set of results has to do with the amounts attached to the historical salaries disclosed. On average, higher salaries increase recruiter beliefs about candidate quality and competing offers. Every extra \$1.00 increase in disclosed salary increases employer WTP by \$0.65, beliefs about the median competing offer by \$0.77, and the salary offer by \$0.68. We find that recruiters *discount* extra dollars given to men through the gender wage gap. An extra dollar given to a male candidate (because of the gender wage gap) raises WTP by only \$0.42 (\$0.48 for the salary offer, and \$0.62 for the median competing offer).

By contrast, recruiters regard an extra \$1.00 coming from working at a high-wage firm, or being well paid within a firm's internal distribution, as being far more informative (each extra dollar increasing WTP by \$0.64–0.70, or \$0.20 per dollar more than the male bonus). These sizes suggest that recruiters *anticipate* overpaid men, interpreting their inflated salaries as being less indicative of value. However, the effects are also significantly above zero; recruiters are discounting the male premium by about half, but far from 100%.

Taken together, our first two results contain common themes. Recruiters believe disclosures both the choice to disclose and the amount—are more informative about a candidate's outside offers than the candidate's underlying quality. In addition, recruiters regard female silence and male salary premiums as less informative. This may reflect awareness of true correlations between gender, compensation, value, and willingness to disclose.

The third set of results from our field experiment details how recruiters synthesize the inferences above into decisions. Effects on salary offers are straightforward: disclosing workers particularly those with high salaries—receive higher salary offers. This is especially true for men, whose disclosures are perceived as more informative. By contrast, we find *negative* effects on who is recommended for a callback. Disclosing workers—especially those with high salaries—are less likely to be recommended at all. Although they enjoy higher salary offers when selected, they are less likely to move forward in the hiring process.

Our results about callbacks go in the opposite direction as those on salary amounts. The reasons for this are theoretically grounded and appear in our data: callback decisions are not a function

of worker quality only; they must also incorporate expected costs. This phenomena arises directly from the empirical results described above: salary history disclosure increases beliefs about outside offers more than candidate quality. At some point, outside offers are so competitive that employers' margins are squeezed. Choices of whom to call back mirror those for our measure of expected employer surplus.

For these reasons, we see that nondisclosers (and low disclosers) are called back more often but are offered lower average salaries. This pattern has gendered consequences: men are less likely to be chosen when they disclose. On the margin, employers interpret higher male salaries as constituting a prohibitively high price tag. Women's disclosures have less of an effect on whether they are recommended for a callback; this is consistent with their disclosure choices being less informative. We find similar results on the amounts: lower amounts disclosed by women increase their odds of being recommended.

As a whole, our results highlight important trade-offs. In our setting, disclosures—and disclosing higher amounts—increase the level of salary offers but decrease callback rates. Normative implications about the ban thus depend partly on how policymakers prioritize these competing effects and the associated risk preferences. Additional compensation is obviously useful; however, additional outside offers can also give workers leverage and options (even at a lower salary rate). They may also allow workers greater choice in finding good matches on nonsalary dimensions.

In our final section, we use the results of our field experiment to evaluate salary history bans and the design choices embedded in these policies. We show that salary-history bans reduce gender inequality in offered wages, conditional on callback. They achieve this in part by reducing the salaries of all workers, and particularly men. The ban has little impact on callback inequality. Some states adopt a "partial" ban that allows employers to ask for salary history after making an initial offer; we show that these design choices are less effective at reducing gender inequality.

This paper provides four main contributions, which we detail in the next section. First, we contribute a novel application of disclosure theory to policies intended to "blind" employers and reshape statistical discrimination, and we field experimental evidence about how disclosures are understood by employers. Second, we provide novel findings about gender differences in job search and how employers anticipate and react to them. We propose a microfoundation for these differences in our setting, and we trace how these foundations affect candidate choices and employer reactions. In several empirical tests, we find that recruiters incorporate expectations about gender differences into their decisions. Although their anticipation does not fully eliminate gender disparities, it does reduce inequalities significantly below their original levels.

Third, we extend the literature on "price as a signal of quality" to a labor-market setting, and we introduce the idea of "price as a signal of competition." Our theory and empirics show how price revelations influence both perceptions of candidate quality and perceptions of outside offers.

In our setting, price is a stronger signal of competition than of candidate quality (particularly for gender-related price differences).

Finally, we develop a new experimental methodology to support the research questions around this topic. Our experimental design requires an extension of the audit methodology we call a "two-sided audit." This gives us a novel, behind-the-scenes look at how salary history disclosures propagate through the hiring process. We gather a rich collection of theoretically motivated outcomes from the field about both wage setting and candidate selection. To our knowledge, this is the first paper to randomize a government-mandated policy in an audit experiment. Our two-sided design creates multiple avenues for studying discrimination more broadly in future research.

The rest of the paper is organized as follows: The remainder of this section describes three related literatures and our contribution to each. Section 2 provides a brief background about the practice of asking job candidates for salary histories. Section 3 describes a theoretical framework of employer updating from salary information, and Section 4 describes our empirical setting. Section 5 lays out our experimental design, and Section 6 proposes specifications. Our experimental results are in Section 7, and we use them to simulate the effects of bans in Section 8. Section 9 concludes with a brief discussion.

1.1 Literature

A large and well-developed literature documents and explores the causes of the income gaps across genders and races.² Our paper is also about the persistence of negative wage shocks in the labor market (Kahn, 2010; Oyer, 2008). Within this larger body, our research is related to several strands of prior research.

Voluntary Disclosure and Statistical Discrimination. Our work relates to policies to "blind" decision makers—for example, by banning employers from seeking candidates' credit scores, criminal records, or drug-test results (Clifford and Shoag, 2016; Agan and Starr, 2018; Doleac and Hansen, 2020; Wozniak, 2015), or through gender-blinding resumes (Behaghel, Crépon and Le Barbanchon, 2015; Åslund and Skans, 2012).

In many settings, this form of blinding requires cooperation from the supply side, which can override blinding through voluntary disclosure. This introduces the potential for unraveling (Viscusi, 1978; Grossman and Hart, 1980; Milgrom, 1981; Grossman, 1981; Jin, Luca and Martin, 2015). In many models, lack of disclosure may be viewed as a negative signal of quality, leading to full revelation. Empirically, in markets with quality disclosures, such unraveling is not

²See Goldin (1990); Solnick (2001); Eckel and Grossman (2001); Blau and Kahn (2006); Thilmany (2000); Blackaby, Booth and Frank (2005); Takahashi and Takahashi (2011); Moss-Racusin et al. (2012); Mazei et al. (2015); Ponthieux and Meurs (2015); Rozada and Yeyati (2018).

always observed (Dranove and Jin, 2010; Mathios, 2000). In surveys, women are less likely to disclose income than men (Goldfarb and Tucker, 2012; Agan, Cowgill and Gee, 2020; Cowgill et al., 2021), and female software engineers are less likely to disclose their skills than are their male counterparts (Murciano-Goroff, 2017). There is a key difference between salary history and other disclosure games: with salary histories, it's possible for disclosures to be "too high" as well as "too low." By contrast, in most disclosure games, the state variable is monotonically related to the principal's objective ("more is always better"). Our conceptual framework demonstrates this potential theoretically, and our data show this phenomenon in a survey and field experiment.

Gender differences in job search. A variety of prior research documents female job candidates as being less aggressive in job search, having lower propensity to enter competitive environments (Niederle and Vesterlund, 2007; Flory, Leibbrandt and List, 2015), self-promoting less (Exley and Kessler, 2019), asking for lower salaries from employers (Roussille, 2020), or being less willing to disclose their salary (Goldfarb and Tucker, 2012; Cowgill et al., 2021).³

In our applied setting, we propose a microfoundation for this difference: the gendered differences in behavior are the byproduct of gendered differences in the psychological costs of disclosure. These cost differences are correlated with gender, but are uncorrelated with latent characteristics such as talent. Although disclosure costs can encompass many things, they are distinct from other theoretical explanations for differences in negotiation behavior—for example, the theories that men enjoy competition more (Niederle and Vesterlund, 2007), that one gender has more biased beliefs about its own abilities (Bordalo et al., 2019), or that genders vary by risk aversion (Croson and Gneezy, 2009; Marianne, 2011; Niederle, 2015). In our framing, the act of disclosing enters workers' utility function directly.⁴

We then draw out the implications for the demand side. We show that anticipating gender differences is useful for the employer, and that treating the same negotiation signals differently by gender is necessary to update beliefs accurately (Spence, 1973; Fryer Jr, 2007). Our model connects gendered differences in negotiation behavior to theories of voluntary disclosure, employer learning, and unraveling.

We then measure these responses empirically in a field experiment. We present novel evidence on the demand side's response to gender differences in negotiating behavior. Our empirical results suggest that corporate recruiters indeed interpret disclosure differently across genders, anticipating less disclosure from women, and punishing them less for silence. We also find that recruiters dis-

³Other examples include the propensity to apply for a job given the number of other applicants (Gee, 2019), the choice to disclose skills (Murciano-Goroff, 2017), the decision to reapply to an employer following a rejection (Brands and Fernandez-Mateo, 2017), and the choice to negotiate wages (Laschever and Babcock, 2003; Biasi and Sarsons, 2020).

⁴Exley, Niederle and Vesterlund (2020) similarly study negotiation costs by gender, including indirect costs of unsuccessful negotiations.

count the higher salaries reported by men. Although recruiters' anticipation is insufficient to fully eliminate gender disparities, it does reduce inequalities significantly below their original levels.⁵

Our paper sheds new light on how employers interpret gendered differences in self-promotion and negotiation aggression. We show that these are not only interpreted as a signal of candidate quality, but also as signals about the competitive landscape. In short, our results suggest that less aggressive job search (such as nondisclosure) affects women twice: once through the employer's own assessment of the candidate's quality, and again through the employer's beliefs about how other employers view the candidate.

Price as a signal of quality. Our paper is also related to prior literature about using price as a signal of quality. This idea was originally applied to consumer products in industrial organization, but is less developed in labor settings. Seminal papers by Wolinsky (1983) and Milgrom and Roberts (1986) study price as a signal of quality,⁶ but do not portray price as a signal of competing offers, possibly because of the presumed thickness/competitiveness of demand for consumer products.⁷

By contrast, labor markets often feature thin and/or monopsonistic demand for workers (Manning, 2003; Ashenfelter, Farber and Ransom, 2010). In this setting, a worker's price can signal not only his or her quality, but the amount and level of competing offers in a worker's search. These beliefs can affect wages through a separate, nonquality channel. We formalize this notion in our model and relate it to common-versus-private-value labor markets. The interaction of these signals is critical to choices about revealing (versus concealing) historical prices.

We then provide direct experimental evidence on how historical prices affect employer beliefs about both quality and outside competition for workers. In most of our results, we find a greater role for price as a signal of competition (versus as a signal of quality). In particular, our recruiters interpret gender-related price differences as signals of competition, rather than of quality.

Audit Methodology. Methodologically, our work is related to recent innovations in correspondence audit methodology (Bartos et al., 2016; Kessler, Low and Sullivan, 2018; Kline and Walters, 2019; Avivi et al., 2021; Cowgill and Perkowski, 2021). A review by Bertrand and Duflo (2017) says, "With a few exceptions, the literature has failed to push the [audit] correspondence methodology to design approaches to more formally test for various theories of why differential treatment is

⁵Murciano-Goroff (2017) says that evidence about anticipation is "lacking" (p. 3). The most closely related papers regarding anticipation are Reuben, Sapienza and Zingales (2014), Exley and Kessler (2019) and Murciano-Goroff (2017). These papers report limited evidence of anticipation using laboratory (Reuben, Sapienza and Zingales, 2014), online subject pools (Exley and Kessler, 2019), and observational (Murciano-Goroff, 2017) designs. Our data do not allow us to compare the degrees of anticipation across these papers.

⁶Roussille (2020) adapts the Wolinsky (1983) model to a labor market.

⁷An unpublished manuscript by Allon, Bassamboo and Randhawa (2012) ("Price as a signal of availability") comes closest to developing this idea in the industrial organization setting of consumer goods.

taking place." In most traditional audit studies, the researcher cannot control the characteristics of the employers. Our two-sided design allows manipulation of employer characteristics in a randomized way. To our knowledge, this is the first paper to randomize a government-mandated policy in an audit experiment. This two-sided design creates multiple avenues for studying discrimination more broadly in future research.

Salary History Bans. A nascent literature directly studies salary history bans introduced in the past decade. The theoretical predictions about the effects of salary history bans are nuanced and ambiguous (Cullen and Pakzad-Hurson, 2021; Meli and Spindler, 2019). A series of empirical papers study salary history bans using panel methods and a variety of observational data sets (Bessen, Meng and Denk, 2020; Davis, Ouimet and Wang, 2020; Hansen and McNichols, 2020; Mask, 2020; Sinha, 2019; Sran, Vetter and Walsh, 2020). A few other researchers have examined the effect of salary disclosures and salary history bans using experiments in online markets (Barach and Horton, 2021), laboratory settings (Khanna, 2020), or in real-life educational institutions (Sherman, Brands and Ku, 2019).

Our field experiment uses recruiters for corporate jobs and is focused on the microeconomics underlying these policies (voluntary disclosure, unraveling, and from prices signaling both quality and competition). Our results address design considerations in policies for blinding decision-making. As the next section shows, there is significant variation in the design of salary history bans across jurisdictions.

Given these research goals, we use an experimental paradigm that delivers access to detailed variables in the hiring function (including WTP, beliefs about competing offers, and expected employer surplus). These variables are typically missing from administrative data sets and do not directly appear in the other experiments, but are useful for understanding the mechanisms and trade-offs around the ban.

2 Background: Salary History Questions

Survey evidence suggests that up to to 43% of job applicants are asked about salary history during job search (Hall and Krueger, 2012; PayScale, 2017; Barach and Horton, 2021; Agan, Cowgill and Gee, 2020; Cowgill et al., 2021). In our own survey, the most common method of inquiring about salary history was on job application forms (in writing). Among workers who were asked, 45% were asked this way (Agan, Cowgill and Gee, 2020; Cowgill et al., 2021). Written salary history questions on job application forms are so common, in fact, that some jurisdictions explicitly

address the practice in the text of their bans.⁸ Job interviews (34%) were the second most common context of the question.

In August 2016, Massachusetts adopted the first ban on salary history questions (effective July 2018), thus becoming the first jurisdiction to adopt such a policy. As of August 2021, 21 states and 21 local jurisdictions have adopted some form of salary history bans.⁹ In 2019, a federal salary-history ban passed the House of Representatives (and passed again in 2021).¹⁰ President Biden agreed to sign it if the federal ban passed the Senate, and has issued an executive order to ban salary-history questions in federal agencies' hiring practices.¹¹ These laws vary in their details but nearly always prohibit oral or written questions about salary history, even if the questions are posed as voluntary or optional.¹²

However, applicants under the bans are still permitted to voluntarily and without prompting disclose salary history information. In most jurisdictions, employers are allowed to use or confirm voluntarily disclosed information.¹³ In addition, the laws do not ban employer questions about a candidate's hopes, wishes, or expectations around compensation.

The popularity of these bans masks enormous heterogeneity in the designs of salary history bans. For example, some bans apply only to either state or jurisdiction agencies (public employers); others apply to all employers in the jurisdiction (both public and private). Our experiment addresses heterogeneity in the design of the ban. Some jurisdictions ban employers from asking until an initial offer has been made (but they can ask afterward during negotiation of the initial offer).¹⁴ This design is focused on giving all candidates a common "floor" to begin negotiating. By

¹¹https://www.law360.com/employment-authority/articles/1401105/biden-targets
-pay-equity-with-salary-history-executive-order

⁸These jurisdictions include the states of Virginia, for public jobs, and New York, as well as municipal-level bans in Atlanta, Ga.; Jackson, Miss.; Suffolk County, N.Y.; and Richland County, S.C. See https://www.hrdive.com/news/salary-history-ban-states-list/516662/.

⁹See https://www.hrdive.com/news/salary-history-ban-states-list/516662/ for the most up-to-date list. This list includes the two states that have banned local jurisdictions from making salary history bans.

¹⁰2019: https://thehill.com/homenews/house/436121-house-passes-paycheck-fairne ss-act; 2021: https://www.cnbc.com/2021/04/16/what-the-paycheck-fairness-act-cou ld-mean-for-women-and-the-pay-gap.html

¹²The goal of the bans is expressed in the title of the laws: "The Act to Establish Pay Equity" in Massachusetts, or "An Act Concerning Pay Equity" in Connecticut. On the day the Illinois law was announced, one of the sponsors said, "I am proud to stand with our new governor today as he takes strong, immediate action to close the gender pay gap and move towards pay equity[.]" https://www2.illinois.gov/Pages/news-item.aspx?Relea seID=19609

¹³Some jurisdictions' bans explicitly grant permission to confirm or use voluntarily supplied salary history information. One exception is California, where employers are expressly prohibited from relying on even voluntarily disclosed information.

¹⁴This design has been adopted in New Jersey, Alabama, Delaware, the District of Columbia, New York (2017–2020, until a revision), and Atlanta. See https://www.hrdive.com/news/salary-history-ban-states-list/516662/ and https://www.ebglaw.com/news/new-york-state-releases-guidance-on-salary-history-ban/.

contrast, other jurisdictions ban asking at any point in hiring, both when making callback decisions and when setting wages. This includes California and New York (from 2020 to the present, following a revision). This approach aims to eliminate salary history not only from the floor, but from the negotiation altogether. This latter distinction about when employers can (if ever) ask about salary history is an important component to understanding the potential impact of salary history bans. It will come up again in our simulations in Section 8.

3 Conceptual Framework

We now offer a simple theoretical framework of hiring and wage setting in which salary history disclosures (or lack thereof) are a key input. This framework will offer insight into when and how salary history disclosures will appear, and how they shape hiring decisions, thus providing guidance for our empirics. Our framework starts from a standard unraveling model (Viscusi, 1978; Grossman and Hart, 1980; Milgrom, 1981; Grossman, 1981), with a few key adaptations to our setting.

Unraveling is a classic result that suggests that if information is verifiable and the costs of disclosure are small, market forces will compel all participants to disclose information voluntarily about their own characteristics. The intuition behind unraveling is simple: An audience of buyers will view all nondisclosing parties as equals, and thus high-quality agents have an opportunity to distinguish themselves through volunteering information (i.e., sharing good news about themselves). When they do, buyers lower their expectations about the quality of the remaining nondisclosers. This logic proceeds iteratively until all parties disclose, or until all nondisclosed salaries can be deduced.

To adapt this model to our setting, we make three adjustments to the baseline model. We summarize these and discuss their implications below, and we use the remainder of this section to specify our implementation of these changes.

- First, the audience for the disclosure (employers) is making two interrelated decisions for each candidate: 1) whom to call back, and 2) how much salary to offer. Consistent with prior research (Postel-Vinay and Robin, 2002; Dey and Flinn, 2005; Cahuc, Postel-Vinay and Robin, 2006), we conceptualize these two decisions as a function of the employer's beliefs about the candidate's value to the employer (WTP), and about the candidate's outside options.
- 2. Second, the audience (employers) does not have preferences directly over the variable that candidates can disclose or not (in this case, salary history). Instead, the disclosed variable is a noisy signal used to update beliefs about the variables above (WTP and outside offer).

Because it is a noisy signal, disclosures may not be fully informative. Below, we discuss some reasons why salary history may be more (or less) informative in labor markets.

3. Finally, all candidates have other observable characteristics (such as gender). Observable characteristics are partially correlated with the hidden variable (salary history). As a result, the employer can make an informed guess about the candidate's hidden value even without a disclosure. This makes disclosure particularly impactful for candidates whose hidden values are rare (given observables).

In addition, candidates have hidden *costs of disclosing*. The presence of disclosure costs means that silent workers contain a mixture of candidates with unattractive salaries (that were strategically withheld) and candidates with high costs. Costs can be correlated with an observable characteristic (in our case, gender). Our survey suggests that women have higher disclosure costs. The implication is that recruiters may know that a silent woman is more likely to have high disclosure costs, and not necessarily a low value, and that a silent man is more likely to have a low value.

Adaptations No. 2 and No. 3 are straightforward. In the remainder of this section, we specify the form of the two decisions (callback and salary level) referenced in No. 1, which the employer makes after receiving the message and updating beliefs. We show that the effects of disclosing, and thus the merits of disclosing, depend partly on whether salary histories contain more information about workers' productivity or about workers' outside options.

3.1 Preliminaries: Players, Utilities, and Notation

The two players in our model are an employer and a job applicant. The employer's utility comes from money and the productivity from hiring a candidate. The employer chooses to call a candidate back (or not) with a take-it-or-leave-it (TIOLI) salary offer. If hired, a job applicant yields a stream of utility payments to the employer over multiple time periods $(v_1 + v_2 + ...)$ and salary payments to the worker $(s_1 + s_2 + ...)$ from the employer. (These payments may be zero after a worker quits or is fired.) The net present values of these streams are v and s. Making a callback costs the employer c.

The applicants' utility comes from salary payments and the disutility of work. Each applicant has a set of k outside offers. The outside offers are the net present value of the salaries offered by other employers. The k offers are drawn from a distribution H, and they include the candidate's current position, if currently employed.¹⁵ A candidate's best outside option, η , is the first-order

¹⁵Additional draws come from a combination of the active search for other employers and from employers' active recruitment of the candidate.

statistic from the k draws from H. k and H are known to the employer, but the realizations are not. For any given salary offer s by the employer, the candidate accepts with probability $Pr(s > \eta)$.

Given a joint distribution of beliefs $F(v, \eta)$ about the candidate's value and outside offers, the employer can calculate a TIOLI offer s^* :

$$s^{\star} =_{s} \mathbb{E} \left[\underbrace{\mathscr{W}(s > \eta)}_{\text{Whether candidate}} \cdot \underbrace{(v - s)}_{\substack{\text{Net value of} \\ \text{employment,} \\ \text{if accepts}}} \right]$$

$$=_{s} \iint \underbrace{\mathscr{W}(s > \eta)}_{\text{Whether candidate}} \cdot \underbrace{(v - s)}_{\substack{\text{Net value of} \\ \text{employment,} \\ \text{probability}}} \cdot \underbrace{f(v, \eta)}_{\substack{\text{Joint} \\ \text{probability}}} dv d\eta$$
(1)

Given this s^* , the employer can then decide whether to extend a callback at all. The employer will extend a callback if:

$$\mathbb{E}\left[\underbrace{\mathscr{H}(s^{\star}>\eta)}_{\substack{\text{Whether candidate accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}}\right] = \iint \underbrace{\mathscr{H}(s^{\star}>\eta)}_{\substack{\text{Whether candidate accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal TIOLI offer s^{\star}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal tioned offer s^{\star}}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal tioned offer s^{\star}}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal tioned offer s^{\star}}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal tioned offer s^{\star}}}} \cdot \underbrace{(v-s^{\star})}_{\substack{\text{Net value of employment, if accepts the optimal tioned offer s^{\star}}}} \cdot \underbrace{(v-s^{\star}$$

Note that higher beliefs about v increase the employer's payoffs from giving a callback and for offering a generous TIOLI amount. However, higher beliefs about outside options increase s^* but decrease returns of sending a callback (unless beliefs about v also increase).¹⁶ The employer cannot justify a callback to an expensive employee unless the employee's contributions make it worthwhile. Neither v nor η can be directly observed. However, the employer can estimate these from observable characteristics, the choice to disclose salary history, and the amount, if disclosed. We now turn to how these beliefs are formulated as a function of disclosures.

3.2 Updating Beliefs Using Salary History

Candidates have a hidden salary variable h that can either be disclosed or not. h can be verified if disclosed, and thus the candidate cannot lie.¹⁷ The candidate's action space is $\{h, \emptyset\}$. Working backwards, we now turn to the candidate's choice to disclose. Candidates receive utility from offers and from paying (or avoiding) a cost to disclose.

¹⁶By higher beliefs about v, we mean a new, joint distribution of beliefs $F(v, \eta)$ in which the marginal distribution of v first-order stochastically dominates the original set of beliefs (and the same for η).

¹⁷This is the standard requirement in the disclosure literature.

If the candidate receives no callback, the payoff from the callback is η (his privately known outside option). If he gets a callback, his payoff is α for receiving a callback at all, and an additional β for every dollar above his outside offer (η). If the amount is below his outside offer, he receives only α . This could be positive because an offer—even at a low salary—can potentially be useful to a job seeker.¹⁸ α and β are known to the employer. The candidate's utility is thus

$$u = \underbrace{\eta}_{\substack{\text{Outside}\\\text{offer}}} + \underbrace{\alpha \cdot Callback}_{\substack{\text{Additional payoff from}\\\text{getting a callback}}} + \underbrace{\beta \cdot \mathscr{W}(s > \eta)(s - \eta)}_{\substack{\text{Additional payoff from getting a callback}\\\text{with a salary above the outside offer}}$$
(3)

Since the employer is making salary choices based on Equation 1 and callback decisions according to Equation 2 (and the candidate knows this), we can rewrite the candidate's utility in terms of the employer's offer s^* , which is in turn a function of the employer's beliefs about the candidate's value and outside offer (Equation 1):

where $\dot{\eta}$ represents the employer's true, privately known value of η .

To understand the candidate's disclosure strategy, we must specify how h is correlated with the candidate's value v and his or her outside offers η . The joint distribution of η , v, and h is publicly known. While this distribution could in theory take any shape, two aspects are important:

Informativeness. v (value) and η (outside option) are correlated with h, and E[v|h] and $E[\eta|h]$ are both weakly increasing in h.

This assumption means that h (the hidden salary variable), is a noisy signal of value (v) and outside option (η) . These correlations could be high in settings such as a common-value labor market, or zero in settings where firms value the same workers very differently—i.e., when there are match-specific sources of productivity (Jovanovic 1979*a*,*b*; Lazear 2009).¹⁹

Monotonicity. The expected value of the worker above his or her outside option $(E[v - \eta|h])$ is weakly monotonic in h. This simply means that, on average, the slope of h on v is steeper than h and η (or vice versa).

In principle, the monotonicity of $E[v - \eta | h]$ mentioned in 3.2 could either be increasing or

¹⁸We have three specific scenarios in mind. First, a low initial salary may still be useful as a bargaining chip in renegotiations with one's current employer. Second, a worker may gain utility from nonpecuniary aspects of the job. Stern (2004) finds that scientists choose a 30% pay cut to be able to do science. Finally, a low starting salary could be compensated for with greater wage growth in the future, either at the same company or by switching. Through any of these mechanisms, a low offer may be useful.

¹⁹In principle, both correlations could be negative for a given employer.

decreasing. High h candidates could, in theory, deliver the highest surplus.²⁰ In this case, as h increases, the worker would be paid an increasingly small fraction of the value he or she provides. This would be true if a high salary offer increased valuations more than outside offers. From the perspective of disclosure strategy, candidates disclosing high amounts receive both higher callback chances and higher salary offers (conditional on a callback) than low disclosers.

However, the relationship could also go in the opposite direction: salary histories could be negatively correlated with employer surplus. High h candidates could deliver lower surplus. As h increases, workers are paid a larger fraction of their value. This would happen if high salaries increase outside options faster than the benefits. At some point, outside options may catch up with value, leaving the candidate too expensive to warrant a callback. In this scenario, high h candidates face a trade-off: disclosing may decrease their callback chances but increase the salary amount conditional on callbacks. Is the trade-off worthwhile? It depends on the value of receiving a callback at a low amount (α).

Which of the two scenarios describes reality is an empirical question we address with our experiment. As we discuss later, our results more closely resemble the second scenario, in which high h candidates are increasingly bad deals: a higher salary history increases beliefs about η more than beliefs about v, and thus at some point, candidates become too expensive to call back. Some candidates thus face a trade-off between high salaries and any kind of callback that depends on α , the utility weight placed on callbacks.

3.3 Choices to Disclose and Unraveling

The candidate can now make a disclosure decision. Disclosure is costly and requires a cost m that is privately known to the candidate and drawn from a distribution F_m that is publicly known. The candidate can anticipate the employer's posterior beliefs about v and η following a disclosure. With these posteriors, the candidate can anticipate whether she will receive a callback from the employer and at what price, using Equations 1 and 2. The candidate can then use Equation 3 to determine the payoff from a disclosure.

If she does not disclose, then the standard unraveling logic proceeds: the employer assumes that the candidate had an unattractive salary and proceeds as if the salary history was bad. This makes disclosure more attractive, even for those with unattractive salaries. As described above, this proceeds until all candidates disclose, or until the benefit of disclosing is equal to the cost, m.

Where our story differs from the standard unraveling story is around what constitutes an "unattractive" salary. Because high h candidates are bad deals, salaries could be "too high" if

²⁰One setting where this appears to be true is professional sports, where leagues with salary caps prevent employers (teams) from paying the full value of superstars. As such, superstars with the maximum permissible contracts under the cap are regarded as cheap compared to the value they bring, and thus a good use of dollars.

candidates place a high weight on getting a salary offer at all (high α). If $\alpha = 0$, the candidate cares only about high salary unconditional on a callback. The strategy is to disclose if high, hide if low. The standard unraveling logic proceeds. Silence is assumed to be a mixture of low salaries and high costs of disclosure (high *m* realizations from F_m). How low a salary should one disclose? Until benefits stop exceeding *m*.

If $\alpha = 1$, the candidate cares only about getting an offer. The best strategy is to disclose if low, hide if high. The standard unraveling logic proceeds *in reverse*: low salaries should be disclosed, and high salaries should be hidden. Silence is assumed to be a mixture of *high* salaries and high costs of disclosure. How high a salary should one disclose? Again, until benefits stop exceeding *m*.

If α is in between, both high and low can be bad. As a result, silence contains high cost candidates and a mixture of high and low *h* candidates. The exact mixture depends on the level of α .

3.4 Observable Characteristics and Costs

As described at the beginning of this section, candidates differ on observable characteristics. Our framework above easily accommodates this; the joint distribution of h, v, and η can differ by observable characteristics. As a result, employers' beliefs and reactions to candidates who behave the same way could differ, even for candidates with an identical h.

We also allow costs m to vary by observable characteristics. Costs of disclosing play a key role in our theory by regulating how far unraveling proceeds, and thus how informative silence is. For candidates with observable characteristics i with high average disclosure costs, unraveling will be more limited. Rather than signaling the worst salary, silent types will contain a mixture of relatively good and bad types (depending on how high costs are).

At first glance, disclosing salary history would appear to be cheap using a pay stub, bank statement, or offer letter. However, workers report feeling *psychological costs* of disclosing. In our survey, a majority of workers expressed that they were uncomfortable answering salary history questions by a potential employer.²¹

In addition, these costs are not evenly distributed. Across multiple studies, women are less comfortable disclosing their salaries than men (Goldfarb and Tucker, 2012; Agan, Cowgill and Gee, 2020; Cowgill et al., 2021). In our own survey, women were about 6 percentage points more

²¹Surveyed workers were asked to rate how much they agreed/disagreed with the statement, "I am fundamentally uncomfortable answering the salary history question when asked by a potential employer," on a seven-point scale from strongly disagree to strongly agree. Answers of five or above were coded as agreeing in some form. More than half, 51%, agreed in some form, with 21% saying they strongly agreed. Some 17% of respondents strongly agreed with the statement, "I can't think of anything that would make me fully comfortable providing my salary history (when asked by a potential employer)," and 44% expressed some form of agreement.

likely to agree with statements about feeling discomfort with disclosing than men. This is salient for our conceptual framework and experiment, because it suggests that a woman's silence will contain less information about their underlying characteristics.

3.5 Implications for the Ban

The ban affects only the employer's ability to ask questions. Most unraveling models do not feature a question, as candidates can voluntarily disclose. We interpret employer questions (or other prompts) as lowering m_i (the costs of disclosing), and thus bans raise this cost for all individuals.

Of course, some candidates will disclose no matter whether the employer asks or not. These are candidates that have relatively attractive salaries and low costs of disclosing. We call these candidates "always disclosers." Similarly, some workers ("never disclosers") will not disclose, even when asked. These are candidates with unattractive salaries and/or high costs. Last, "compliers" will disclose when asked, but will be silent when not asked. Compliers' salary histories are neither exceptionally attractive nor unattractive, and thus their disclosure costs are pivotal.

We summarize this typology in the 2×2 matrix in Figure 1. In principle, a fourth type exists: "Defiers," who are silent when asked and volunteer when not asked. However, our theory suggests why this group would not exist: if the ban raises costs of disclosing, all candidates should be less likely to disclose, not more. When we investigate the existence of defiers in our survey data, we find they are less than 0.5% of the U.S. workforce.

4 Setting

To apply our conceptual framework to data, we examined disclosure in the labor market for software engineers. We focus on this industry for reasons outlined in Appendix A: broadly speaking, the characteristics of this industry help place our results in the context of wider labor market trends. Our experimental design utilizes the managerial practice of delegating recruitment decisions to recruitment specialists, whom we hire to make decisions about candidates in our two-sided audit study. In this section, we provide an overview of this setting and review institutional details that motivate our experimental design.

4.1 Outsourced Recruiting

In the past two decades, the delegation of recruiting has become widespread (Cowgill and Perkowski, 2021). Although comprehensive data on firm hiring practices is not available (Oyer, Schaefer et al., 2011), surveys estimate that between 40% and 50% of U.S. firms outsource some portion of their

recruiting process or plan to do so soon, and that as many as one-third of jobs in the United States are filled through outsourced recruiting.²² Firms either hire individual recruiters on a temporary contract basis, or they outsource recruiting work entirely to a third-party organization. A large industry known as the recruitment process outsourcing (RPO) industry exists to support this practice. The recruitment process outsourcing industry has been growing steadily since 2000, and is growing faster than other sectors.²³. The subjects in our field experiment are recruiters employed in this industry.

4.2 Recruiter Work

Industry benchmarking reports show that recruiters perform a wide variety of tasks for employers.²⁴ Surveys indicate that nearly all firms that use outsourced recruitment ask their recruiters to screen applications, and over 95% of recruiters have been asked to provide input on the salary of potential hires.²⁵ Even before the COVID-19 pandemic, over 80% of outsourced recruiting was performed remotely.²⁶ Recruiters in this industry are typically told to avoid searching on the Internet for information on job candidates, as this can violate employment law.²⁷ Our study design closely mimics each of these attributes.

4.2.1 Structured Evaluations

Recent surveys of recruiters by Jobvite,²⁸ Monster.com,²⁹ and Black, Hasan and Koning (2020) indicate that assessing candidates using structured and/or quantitative evaluation is typical and

²²https://www2.staffingindustry.com/content/download/272268/9900694/Staffing Trends_Free_190301.pdf https://www.workforce.com/2019/01/24/recruitment-proce ss-outsourcing-providers-think/.

²³According to Cowgill and Perkowski (2021), outsourced recruiting has been growing at 4 to 10 times the rate of overall U.S. employment since 2000, using the BLS's Occupational Employment Statistics (https://www.bls.gov/oes/home.htm) Additionally, the number of recruiters in employment services has more than doubled, while U.S. employment overall grew only 15% from 2002 to 2009, according to the U.S. Economic Census. The growth in outsourcing recruiting practices is driven by many factors, including increases in the popularity of Internet job search, numbers of applicants, demand for employee screening (Autor, 2001; Cappelli, Wilk et al., 1997), and returns to selective hiring (Acemoglu, 1999, 2002; Levy and Murnane, 1996)

²⁴See https://www.shrm.org/ResourcesAndTools/business-solutions/Documents/Ta lent-Acquisition-Report-All-Industries-All-FTEs.pdf for example.

²⁵See https://staging.kornferry.com/media/sidebar_downloads/Measuring-Up-A-ne w-research-report-about-RPO-metrics.pdf and Analysts (2017), as well as our own survey responses from subjects in this experiment.

²⁶Staffing Industry Analysts, RPO Market Developments, December 2017.

²⁷For example, the Equal Employment Opportunity Commission (EEOC) tells firms to avoid online searching for candidates. See https://www.lexology.com/library/detail.aspx?g=1147c039-ef9c-4f6a-9ebb-448de20b8123.

²⁸https://www.jobvite.com/wp-content/uploads/2015/09/jobvite_recruiter_natio n_2015.pdf

²⁹https://www.monster.com/about/a/monster-2018-state-of-recruiting-survey

expected to grow. An industry organization, the Talent Board, found that 71% of employers used preemployment assessments and selection tests in 2019.³⁰ Westin (1988) reviews the history of quantitative assessment and formal job testing. Employment testing grew out of twentieth-century personality psychology, which has a long tradition of quantitative assessment scales.

4.3 **Recruiter Incentives**

The use of monetary incentives is widespread in recruiting. The monetary incentives a recruiter faces can take many forms and are often quite complex.³¹ Employers often use a formula for bonuses that rewards recruiters for selecting candidates who are eventually hired and rated as high performers (either on the job or during the selection process).³²

If recruiters were rewarded for selecting candidates who are hired, they may be tempted to forward candidates indiscriminately. Recruiter incentives therefore can include measures to induce selectivity, so that the employer does not waste time on candidates unlikely to match. In some cases, firms impose an explicit cap on the number of candidates a recruiter can forward. In settings where an employer is seeking all qualified candidates, firms may use an explicit monetary penalty for forwarding candidates who are a bad fit.³³ Beyond monetary incentives, recruiters face reputational incentives. Even without explicit monetary rewards or penalties, recruiters know an employer wants them to forward a curated list of candidates and give useful advice about fit and compensation. Explicit monetary schemes reinforce these implicit reputational concerns.

³⁰https://www.thetalentboard.org/report/the-2019-talent-board-north-american -candidate-experience-benchmark-research-report-now-available/

³¹Recruiter.com states that the "complexity of recruiter commission plans tends to rival both ontological arguments and mortgage refinancing documents." See https://www.recruiter.com/i/recruiter-commission -plans/

³²According to industry reports, 60% of performance pay is measured by the number of candidates hired. See https://www.shrm.org/resourcesandtools/hr-topics/talent-acquisition/pages/rewarding-recruiters-for-performance.aspx. A particularly well-documented example of this practice is the U.S. military's recruiter bonus programs (Condren, 1997; Asch, 1990). The Bureau of Labor Statistics' National Compensation Survey (NCS) found that 43% of human resource specialists received performance pay in the first quarter of 2020 (Makridis and Gittleman, 2020). This percentage likely understates the extent of incentives for recruiters; the NCS definitions of performance pay include incentive payments that "follow an explicit formula" and discretionary spot bonuses (Gittleman and Pierce, 2015). Other sources of performance-based rewards that are excluded include permanent wage increases, promotions, and the possibility of new assignments or repeated business (for contracted recruiters).

³³HR Magazine https://www.shrm.org/hr-today/news/hr-magazine/pages/1103hirschm an.aspx

5 Experimental Design

To examine the impact of salary disclosure on candidate outcomes, we implemented a two-sided audit study. We hired a real recruiting workforce to screen (fictitious) candidates on behalf of a (fictitious) firm. This design allows us to vary both characteristics of the candidates *and* characteristics of the firm. To our knowledge, this design has been utilized in only one other paper (Cowgill and Perkowski, 2021), and we extend the design to fit our research agenda. On the candidate side, variation comes from experimental changes to candidates' gender, salary, and whether that salary is disclosed at all during the application process. On the employer side, there is variation in whether the firm asks for salary information, and the number of candidates who disclose their salary history. Each recruiter's job was to review eight candidates and provide feedback about each to the firm, with questions and format detailed below.

5.1 Our Recruiting Workforce

We hired 256 recruiters to evaluate 2,048 job applications for our field experiment. These recruiters are our experimental subjects. We identified recruiters who are typical of those hired by companies through the recruitment outsourcing industry and engaged them in the usual way for this industry. Using a large and popular outsourcing platform, we identified and contacted professional recruiters as discussed in Appendix C. We only contacted recruiters who had prior recruiting experience and a U.S.-based location. We offered to pay recruiters their hourly rate as it appeared on their profile, with an additional bonus, to be described shortly.

The human resource workers in our field experiment were similar to those in the U.S. as a whole.³⁴ Each subject was assigned to one of the experimental conditions (described below) using the randomization procedure in Appendix B.1.

5.2 Our Recruiting Task

Recruiters who accepted our offer were given three documents to complete their task: 1) detailed instructions describing the job they were hiring for, 2) eight one-page PDF job-candidate applications, and 3) an online structured evaluation form to provide feedback about the candidates. Examples of these are provided in Appendix H.

 $^{^{34}}$ According to the BLS, in 2018, human resource workers across all industries were 69.7% female and 10.5% black. The median hourly wage was \$29.01 across all industries and \$41.93 in the software industry.³⁵ As compared with the BLS statistics about human resource workers in the U.S., the recruiters in our study were slightly more likely to be female (75%), twice as likely to be black (23%), and had a higher hourly wage of \$44 (Table G2). The BLS does not report demographic characteristics of industry × occupation cells. However, these figures can be calculated using the Five-Year (2012–2017) American Community Survey Public Use Microdata. There are approximately 115 human resource specialists in the software industry in this sample. They are approximately 80% female, 75% white.

As described in Section 4.2, there is wide variation in the tasks recruiters perform. Our goal was to design a task and incentive scheme that allowed us to answer our research questions while fitting naturally into our setting. In addition to their hourly rate, we offered recruiters an additional bonus, described in Appendix C. In the main text of our communications with recruiters, we described in simple, nontechnical language how the bonus worked. This was likely sufficient for many of our recruiters. However, the details of the bonus, including an explicit formula, were available in the FAQ portion of the same PDF document.

Recruiters were required to answer each question about all candidates before proceeding to the next question for any candidate. Recruiters were not told the specific items on the structured evaluation beforehand. Each group of questions was on a separate page, which did not allow recruiters to revise previous answers.³⁶

Below, we detail the items in the structured candidate evaluation. In this paper, we describe recruiter choices in this evaluation using theoretical terminology. However, these concepts also appear in typical hiring settings in less formal language, without models or academic jargon. Recruiters and hiring managers often consider questions like, "How much will this worker contribute?" "How much are we willing to pay?" or "What will their competing offers look like?" Our experiment measures these common economics concepts without the use of economics terminology. A full copy of the structured evaluation can be seen in Appendix H.6.

5.2.1 Primary/Incentivized Assessments

The three variables below (callback, offer, and WTP) were the first items evaluated by recruiters. These items were also explicitly used in each recruiter's bonus formula (and many counterpart bonus formulae in the real world).³⁷ The recruiters therefore faced monetary incentives to set these numbers correctly.

Callback. Like a traditional audit study, we observe whether each recruiter recommends a candidate for a callback. Recruiters were aware that they should feel free to suggest multiple callbacks (as many as they deemed a good fit) rather than limit themselves to filling a single position. This is common in high-tech labor settings featuring high demand for qualified workers. We conceptualize this as the callback decision modeled in Equation 2 of our framework.

Salary Offers. Recruiters made a take-it-or-leave-it (TIOLI) salary offer for each job candidate. We conceptualize these as the s^* choice modeled in Equation 1 of our framework. Hall and

³⁶If we later found an error in an item (e.g., a recruiter typed in letters in an item that required numbers), we then contacted the recruiter and sent that recruiter a new link to revise that item.

³⁷See Appendix H.5 for details of our specific bonus formula and how it was communicated to recruiters.

Krueger (2012) find that two-thirds of workers report believing that the offer they were made by an employer was a take-it-or-leave-it-offer. We observe a TIOLI offer even if the recruiter did not believe the candidate should be called back.³⁸ The offer and callback decisions were on the first page of our evaluation. After these responses were submitted, recruiters could not revise them.

Willingness to Pay (WTP). According to Barach and Horton's 2021 survey, employers make the first offer about 60% of the time. In the remaining 40% of cases, a job candidate makes a first offer. As a result, our recruiters reported the maximum offer from the candidate that the firm should accept.³⁹ We conceptualize these as the v (expected value) that appears in our framework. By reporting a threshold, we essentially observe the recruiter's value for the candidate directly. Any value below the candidate's true value is a dominated strategy (Becker, DeGroot and Marschak 1964).

5.2.2 Additional Assessments

After the questions above, recruiters made additional assessments that were not explicitly mentioned in the bonus criteria. Although recruiters did not face formal contractual incentives to formulate these numbers precisely, they still face reputational incentives for accurate answers.

Outside Offer Distributions. For each candidate, recruiters state TIOLI salary offers that the candidate would be highly likely to accept, highly *unlikely* to accept, and indifferent about accepting. These were specifically defined to mean salary offers the candidate would accept with 95% probability, 5% probability, and 50% probability. We interpret these as the recruiter's beliefs about the distribution of the candidate's best outside offer—in our conceptual framework, η —at the 5th, 50th, and 95th percentiles of this distribution.

Competing Offers. Recruiters estimate how many competing offers each candidate would receive during his or her search from other employers. To simplify this task, recruiters could choose either "zero or one" or "two or more." In our conceptual framework, this roughly corresponds to k, the number of outside offers drawn from the candidate's outside-offer distribution. Prior research suggests that outside offers increase the bargaining power of the candidate (Blackaby, Booth and

³⁸Recruiters made recommendations about the annual base salary of compensation only, although the firm instructions said, "We also offer benefits including health insurance, stock, and performance-based annual bonuses," without specifying their amounts. To observe salary offers even for those not suggested for a callback, recruiters were told, "For candidates you do NOT suggest interviewing, please enter the amount you think they should be offered were they to pass an interview—this may be helpful for us in the future."

³⁹Recruiters could choose a maximum between \$20,000 to \$200,000 in \$10,000 increments. One can think of this as asking, "Should the firm be willing to pay \$20,000, \$30,000, etc., up to \$200,000?" Recruiters did not appear to feel constrained by this range, since the minimum valuation was \$60,000 and the maximum valuation was \$180,000.

Frank, 2005), and that employed workers rarely receive more than one job offer at a time when searching.⁴⁰ Recruiters also state whether competing job offers from each candidate would come from the candidate's own search efforts or from rival employers' search efforts. These variables shed light on why the levels of competing offers (η) are as high or low as they are.

5.2.3 Composite Outcomes

We take some of the main items we observe from recruiters and combine them into the following outcomes of interest.

Surplus. We label the value of the candidate as measured by willingness to pay (WTP) minus the suggested salary offer as the surplus a recruiter assigns to a candidate. Our theoretical framework, Section 3, describes how this can be interpreted as the margin the employer would get from this candidate. This corresponds to $E[v - s^*]$ in our conceptual framework, and it appears in Equations 2 and 4. It is also the topic of the monotonicity requirement in Section 3.2 (Assumption 3.2), which states that surplus is an increasing or decreasing function of salary history. As discussed in 3.2, the direction of this monotonicity determines whether high-salary history candidates face a trade-off between receiving a callback at all (and sacrificing the salary amount) versus preferring a higher amount (and a lower chance of receiving a callback at all).

Outside Offer Range. We define the outside offer range as the difference between the 95th percentile and the 5th percentile of the outside offer distribution. This gives us information about how wide-ranging the recruiter's beliefs are about a candidate's outside offers. Larger ranges indicate more diffuse beliefs. This is akin to the variance of the belief distribution of η (the best outside offer).

Probability of Accepting TIOLI Offer. This variable represents how likely the recruiter believes the candidate is to accept the take-it-or-leave-it salary proposed by the recruiter at the start of the form (described in 5.2.1). In our conceptual framework, it is represented as $Pr(s^* > \eta)$ and appears in Equations 1, 2, and 4. To estimate this probability, we fit a logistic curve through the reported 5th, 50th and 95th percentiles of the outside-offer distribution.⁴¹ We used this fitted model to predict where the salary offer made by the recruiter falls on that curve. In our theoretical framework from Section 3, we discuss how the probability of accepting an offer relates to whether the salary offer is above or below a candidate's outside offer.

 $^{^{40}}$ Faberman et al. (2017) find that only 29.1% of employed workers who are looking for work receive at least one offer per month.

⁴¹Similar exercises using probit curves produced results with correlations of 0.99 with the logistic approach.

Probability Accept × **Surplus.** The expected surplus for the employer can be thought of as the probability a specific salary offer is accepted (as described above), multiplied by the surplus should that take place. In our conceptual framework, it is represented as $Pr(s^* > \eta)(v - s^*)$ and appears in Equations 1, 2, and 4.

5.3 Experimental Manipulation

Having described our outcomes, we now turn to experimental manipulations. Our two-sided audit study entails experimental manipulation on both the employer side *and* the candidate side.

5.3.1 Employer Side

On the employer side, we manipulated whether the job applications reviewed by a recruiter asked the candidates for their previous salary, or not. In our survey of workers, we found that 44.8% of workers who were asked about their previous salary were asked on the job application (Cowgill et al., 2021). Asking about salary history on the job application form is so common that several states specifically ban salary-history questions on the job application form (in addition to verbal questions). Our question asked for the applicant's annual base salary at his or her current or most recent job.

5.3.2 Candidate Side

In addition, we randomized the candidate's answers listed on the job form. Our candidate randomizations fall into two categories. The first is related to the candidate's biographical details, and the second is related to the candidate's disclosures.

Biographical Details. As stated in our theory section, we had two hypotheses about biographical details. The first is about gender. To randomize gender, we created candidate names using the top four male and the top four female names from American cohorts of 1991–1994 according to the Social Security Administration, making the candidates a few years out of college at the time of our experiment.⁴² Our hypotheses about gender are also related to the disclosed salary amounts (described next).

The second randomized biographical detail is about low- and high-wage firms. To randomize this, we utilized a list of the 13 biggest employers of software engineers from Monster.com and

⁴²The names were, for the boys, Andrew, Christopher, Joshua, and Tyler; and, for the girls, Emily, Jessica, Samantha, and Sarah.

Indeed.⁴³ This included substantial variation in median salaries for software engineers who had recently graduated.⁴⁴

In order to present job applications to recruiters, we needed additional characteristics for candidates besides a first name and a former employer. Our goal with these other characteristics was to hold them roughly constant at values representative of the broader software market. Some details were held constant: all candidates held a bachelor's degree in computer science, and none required a work visa. However, we permitted some additional random variation in other biographical details in order to avoid suspiciously identical candidates. Appendix D lists these details.

Using this procedure, we created 32 candidate biographies divided into four packets of eight candidates. Each packet contained four male and four female candidates, with randomly chosen former employers. Each packet was then assigned to a treatment and subtreatment condition, described below. Thus, these 32 candidates were evaluated under different experimental circumstances (one packet per recruiter). By asking recruiters to evaluate the same 32 candidates, our experiment permits "biography fixed effects" to increase the efficiency of the experiment.⁴⁵

Salary History Disclosures. Candidates' salary history disclosures (or lack thereof) were also randomized. In packets where candidates were asked salary histories, candidates who disclosed that information completed the question on the form by entering a number on the line. In packets where candidates were *not* asked, candidates who disclosed it used an optional field for "Additional Skills and Information."⁴⁶

The amount of disclosed salary is also randomized (among those who disclosed). We designed these disclosures to be consistent with the candidate's biography but to include some random variation (conditional on the biography). For each candidate's current employer, we looked up the distribution of salaries for software engineers at the candidate's location and job level using Payscale.com.⁴⁷ Candidates were assigned a salary that was either relatively high (near the 75th

⁴³https://www.monster.com/career-advice/article/top-tech-employers-job-listi ngs and https://www.techrepublic.com/article/the-10-companies-hiring-more-soft ware-engineers-than-anyone-else-in-silicon-valley/.

⁴⁴According to PayScale.com, the highest was Oracle (median salary of \$126K), and the lowest was General Dynamics (median salary of \$73K).

⁴⁵A biography consists of the specific combination of name, gender, previous and current employers, job titles and descriptions, dates of employment, undergraduate education, and skills.

⁴⁶This section also included information about programming skills the applicant had. For disclosure without a prompt, the candidate would add a sentence such as "Current base salary: \$X yearly." We randomized whether the statement appeared before or after the additional skills, as well as varying the language of this statement slightly.

⁴⁷Data from websites like Payscale.com and Glassdoor.com are self-reported by workers who visit these websites, which could mean these data are inaccurate. However, Glassdoor has periodically compared its data to that provided by the census, and they've found that the distribution of base salaries reported is very similar (Glassdoor, 2019). We used data from Payscale.com because it offers the ability to see the distribution of salaries by company, job roles, city, and level of experience (a level of granularity that is not publicly available from sources like the CPS, BLS, census, or Glassdoor.com).

percentile) or relatively low (near the 25th percentile) within their current firm's salary distribution.

Importantly, we also built in a gender wage gap. These gaps were based on real-world gaps in this industry. Female candidates were given a salary that was 85% of the male salary at the same company/location.⁴⁸ Despite the gender wage gap built in among our candidates, most variation in the disclosed wages can be explained by previous employers and 75th/25th fixed effects. As such, we present targeted, disaggregated specifications in our results that isolate the effects of gender-related wage differences. Appendix D contains additional details about how we assigned salaries to candidates.

5.4 Subtreatments

In theory, all of the attributes above could be randomized independently at the candidate level. However, we clustered randomization in three ways. First, we held the employer features constant within each pack of eight applications. Recruiters were hired to screen applications from one firm, and thus we kept the application materials consistent within the recruiter's packet. Either all eight applications asked for salary history, or all eight did not.

Second, we clustered candidates' disclosure choices. Each packet contained either zero, four, or eight candidates disclosing. When four candidates disclosed, we randomized which four disclosed and sent a separate packet to another recruiter, flipping the candidates' disclosures. This form of clustering allowed us to measure how candidates' disclosure choices affected each other through spillovers.

Finally, we clustered candidates' disclosure amounts. As previously mentioned, candidates' disclosure amounts were randomized (for those who disclosed). We clustered these disclosure amounts so that in some cases, more disclosing candidates were on the high end of their previous employer. This allowed us to measure and control for potential spillover in the monetary amount disclosed (i.e., how one candidate's high disclosures affect other candidates' outcomes).

For the full details of our clustered randomization scheme, see Appendix B. The clustered randomization produced 22 subtreatments, where a subtreatment is a combination of {asked, not} \times {all disclose, half disclose, other half disclose, none disclose } \times {all high amounts, half high + half low, other half high + half low, all low amounts}.⁴⁹ In some specifications below, we control for subtreatment fixed effects.

⁴⁸To approximate realistic gender gaps in salaries, we analyzed data from the 2015 American Community Survey (ACS) on individuals in computing jobs and found women earned around 85% of men in this data source. Publicly available salary data about specific firms—including the sources we used above (Glassdoor and PayScale) and all others we consulted—do not contain gender-specific wage values.

⁴⁹Our total number of treatments is less than $2 \times 4 \times 4 = 32$ because in cases where no candidates disclose, amounts are irrelevant.

5.5 Randomization Procedure and Balance

Because we randomize both sides of the market, we check for randomization balance in both candidate and recruiter characteristics. Appendix D.1 and B.1 show our candidate- and employer-prompting manipulations were uncorrelated with each other (partly by construction) or with the characteristics of the assigned recruiter. Appendix B contains the details of how recruiters were assigned to treatments.

6 Specifications

Our goal is to estimate how candidate salary disclosures—both their existence and amounts—affect hiring outcomes. Our specifications come mostly in one of three forms, outlined by the equations below.

6.1 Specification 1: Effects of Disclosure vs. Silence

 $y_{i,j} = \beta_1 SalaryDisclosed_{i,j} + \beta_2 SalaryHistoryAsked_j + \nu_i + \beta_3 [RecruiterControls_j] + \epsilon_j$ (5)

where *i* indexes candidates and *j* indexes recruiters. Outcomes $y_{i,j}$ are the assessments given to candidate *i* from recruiter *j*. Whether the recruiter saw applications with a salary history question is controlled for by $SalaryHistoryAsked_j$. Our bundle of recruiter controls (*RecruiterControls*_{*i*,*j*}) includes the gender, race, experience level, and hourly rate of the assigned recruiter (which were balanced by design, section B.1). ν_i signifies candidate biography fixed effects (see section 5.3.2).

We also extend this specification in two main ways: First, we introduce candidate gender interactions with the $SalaryDisclosed_{i,j}$ terms. The main effect of candidates' gender is absorbed by the biography fixed effects (ν_i). Second, we replace $SalaryDisclosed_{i,j}$ with a vector { $DisclosedLowSalary_{i,j}$, $DisclosedHighSalary_{i,j}$ }, where $DisclosedLowSalary_{i,j}$ is a dummy variable for whether the applicant disclosed a salary at the 25th percentile of his or her current firm salary distribution, and $DisclosedLowSalary_{i,j}$ is whether that applicant disclosed at the 75th percentile of his or her current firm salary distribution.

6.1.1 Interpretation

We interpret the $SalaryDisclosed_{i,j}$ coefficient (β_1) in Equation 5 as the average effect of disclosing. This represents whether disclosers are (on average) paid more or less than silent types. A positive coefficient indicates that recruiters believe that above-average types have selected into disclosing. When we include, instead, whether the candidate disclosed low (25th percentile of within-firm salary distribution) or high (75th percentile of within-firm salary distribution), this helps identify where in the distribution silent candidates are presumed to lie.

6.2 Specification 2: Effects of Disclosure Amounts

The next specification measures how the amount the candidate discloses impacts outcomes by adding $AmountDisclosed_{i,j}$ terms to Equation 5.

$$y_{i,j} = \beta_1 SalaryDisclosed_{i,j} + \beta_2 SalaryDisclosed_{i,j} \times AmountDisclosed_{i,j} + \beta_3 SalaryHistoryAsked_j + \nu_i + \beta_4 [SpilloverControls_{i,j}] + \beta_5 [RecruiterControls_j] + \epsilon_j$$
(6)

We set AmountDisclosed = 0 for candidates that did not disclose (their overall impacts are captured by "SalaryDisclosed = 0"). As described in Section 5.4, our experiment was designed to study potential spillovers between candidates' disclosures. Our regressions thus include a set of SpilloverControls_{i,j}. We treat these as control variables and do not report spillover coefficients in our tables in this paper.⁵⁰ Spillover controls are detailed in Appendix E.

6.2.1 Interpretation

The slope coefficient on $AmountDisclosed_{i,j}$ (β_2) represents how much outcomes $y_{i,j}$ are influenced by the amount disclosed. Coefficients with larger magnitude (steeper slopes) indicate greater influence on a recruiter's evaluation. Greater magnitudes are consistent with recruiters relying more on the disclosure amount—i.e., updating or changing their beliefs about the candidate based on the information in the disclosure amount. In the opposite extreme (in which the $AmountDisclosed_{i,j}$ slope is flat), recruiters do not incorporate new information from the disclosure amount, and recruiter behavior does *not* change at all as the amount varies. We will use this specification to measure the direction of our monotonicity assumption (#2 in Section 3.2). This requires that the employer's surplus be increasing (or decreasing) in salary-history amounts, or that the salary history has a greater slope with valuation than outside offers (or vice versa).

6.3 Specification 3: The Heterogeneous Effects of +\$1 (or Components of Disclosure Amounts)

Finally, our results also include a third specification that focuses on the subcomponents of $AmountDisclosed_{i,j}$. Here, we unpack the reasons why candidates in our experiment are paid differently, and we show

⁵⁰Our main interest in this paper is the average direct effects of one's disclGeeosures on one's own outcomes.

how recruiters' reactions depend on these differences. As discussed in our candidate variation section (5.3.2), candidates are paid differently for three reasons: 1) Some work at higher- or lower-wage firms; 2) some are relatively well or poorly paid within their firm's distribution; and 3) some are male or female, and thus report differently even conditional on other factors.

A plausible null hypothesis is that recruiters will treat each source of pay variation equally. Each candidate presents a single number (the sum of these factors), not individual components. Nothing in our experiment alerts recruiters to *any* possible reason for pay variation, much less the three above. Nonetheless, recruiters could possibly anticipate these reasons. Certain employers are well known for paying well. Recruiters who know a candidate works for Apple may adjust expectations upwards about prior salary. Recruiters who know that women are typically paid less may adjust expectations about salary history downward. As discussed earlier, the only source of variation that cannot be anticipated is the within-firm variation, which we designed to be uncorrelated with any observable feature.

In order to measure these effects in our data, we utilize the following identity:

$$AmountDisclosed_{i,j} = OverallAverage + FirmOffset_i + GenderOffset_i + WithinFirmOffset_{i,j}$$

$$(7)$$

Each candidate's disclosure amount is the sum of an overall average across all candidates (*OverallAverage*, a constant), plus a firm-specific offset for the candidate's employer (*FirmOffset*_i: some firms pay higher or lower to everyone on average), a gender offset (*GenderOffset*_i, penalizing women and favoring men), and a within-firm offset (*WithinFirmOffset*_{i,j}, representing pay variation for the same job). The sum of these is (by definition) equal to the total amount disclosed. These relationships flow directly from our procedure for creating salary disclosure amounts (Section 5.3.2).

We use this definition to replace $AmountDisclosed_{i,j}$ in Equation 6 with the subcomponents in Equation 7. This leads to our third specification:

$$y_{i,j} = \beta_1 SalaryDisclosed_{i,j} + \beta_2 SalaryDisclosed_{i,j} \times FirmOffSet_i + \beta_3 SalaryDisclosed_{i,j} \times GenderOffset_i + \beta_4 SalaryDisclosed_{i,j} \times WithinFirmOffset_{i,j} + \beta_5 SalaryHistoryAsked_j + \nu_i + \beta_6 [SpilloverControls_{i,j}] + \beta_7 [RecruiterControls_j] + \epsilon_j$$

$$(8)$$

This decomposed regression allows us to obtain separate coefficients for each source of variation in salary and compare them. We use these coefficients to study how recruiters evaluate an extra dollar of historical salary differently, depending on *why* the salary is higher or lower. Earning +\$1 extra because a worker is at a high-wage firm may evoke a different response than if the same +\$1 came from being a star engineer at a lower-wage firm. The extra +\$1 may evoke a different response if it was awarded for being male.

Because prior employer, gender, and within-firm salary are randomly assigned, these separate coefficients can each be interpreted causally. We call these the heterogeneous effects of disclosing an extra \$1. We can separate these effects from the direct effects of being male and employed at a high-wage firm, because the extra +\$1 effects are only present in our observations where candidates disclose.

As with the slopes in Equation 6, the magnitude of these coefficients represents how informative they are. If recruiters anticipate any differences by using biographical features (Apple and/or male workers are typically paid more), this anticipation would *reduce* the informational content of the disclosure and push these coefficients toward zero. Of course, gender and employer differences may *not* be fully anticipated, and thus we could also find nonzero results.

7 Results

By design, our experiment collected information about not only final outcomes (such as whom to make an offer to and what the salary amount offered should be), but also on the inputs into those decisions, such as the maximum salary offer the firm should accept (willingness-to-pay) and information about employer perceptions of the candidate's outside offers. We start by showing how these measures of inputs and outcomes are related in ways our conceptual framework portrayed. We then examine how a candidate's silence or disclosure impacts employer perceptions about these inputs, and how these change with the amount disclosed. Finally, we explore the implications for callback and salary decisions.

In Table G3, we show correlations between outcomes and recruiters' WTP for candidates and their beliefs about competing offers—relationships in our data modeled by our theoretical framework. As expected, recruiters' TIOLI offers to candidates are positively correlated both with willingness-to-pay for the candidate, and with recruiters' beliefs about the candidates' competing offers. The TIOLI/WTP relationship is stronger than the relationship between TOILI and outside offers. However, as our theoretical framework showed, callback decisions can be positively correlated with willingness-to-pay, but negatively correlated with competing offers. Higher outside offers increase the price that firms must pay to attract a candidate and thus reduce the firm's surplus for selecting the worker. Although these results are descriptive, they are in line with our portrayal in the conceptual model.

7.1 Effects of Disclosure vs. Silence

7.1.1 Overall / Average Effects

Our recruiters believe nondisclosers have lower quality (as measured by WTP) and lower outside options, and are given lower salary offers compared with the average disclosing worker. This pattern can be seen in our basic descriptive statistics in Table 1: recruiters' WTP, assumptions about outside offers, and offers are lower for candidates who do not disclose—column (3)—compared to those who do—column (2).

In the odd columns of Table 2, we model these outcomes using Equation 5. When a candidate discloses, recruiters increase willingness-to-pay by about \$6,800 (6.46% over the mean of nondisclosers). Likewise, disclosure increases perceptions of outside offers by \$8,400, an 8.8% increase over the mean for nondisclosers. These results suggest a level of sophistication by recruiters, who appear to anticipate positive selection into disclosing. Disclosure choices have a larger effect on recruiters' beliefs about candidates' competing offers, and a smaller effect on recruiters' willingness to pay—but both are affected in the same direction. These increases in WTP and outside offers lead to an increase in offers of about \$7,300, or a 7.5% increase, in the amount offered to those who disclose their previous salary. Our results in Table 2 suggest that recruiters make inferences about and assign outcomes to silent workers that are worse than the average candidate within the candidate's {current employer \times job title \times gender} wage distribution. In column (7), we see that disclosure, perhaps unsurprisingly, reduces uncertainty about the level of outside offers—recruiters' perceptions about the spread between an offer a candidate would accept with a 5% probability and one accepted with a 95% probability are compressed by nearly \$9,000 (32% over mean spread of \$28,500).

The results above focus on the average effect of silence versus disclosing, but our design allows us to decompose the effect into low versus high earners within each firm. In Table 3, we show results from our extension of Equation 5, splitting Salary Disclosure into disclosing a low (within current firm) salary versus a high (within current firm) salary. We find that recruiters infer that silent candidates' hidden salaries are at or just slightly below the 25th percentile of potential outcomes (given the candidates' observables). Workers below this percentile are actually better off staying silent. However, despite the negative inferences we document, we also see that recruiters do not fully punish nondisclosure as much as would be possible. In fact, some workers who disclose low amounts are regarded as even worse than nondisclosers (they might have been better off staying silent).

"Below the 25th percentile" is a significant discount. However, the theoretical literature on disclosure rationalizes much more punitive discounting: silent workers should be assumed to have the worst possible wage, given observables. Why aren't recruiters more punitive? As in other

settings (Jin, Luca and Martin, 2015), recruiters may be inattentive or naive about the strategic aspect of nondisclosure.

We find evidence of two additional mechanisms. The first is that salary histories can be "too high" as well as "too low." As discussed previously, this is a key differentiator between salary history and other disclosure games in the literature. How can salary histories be "too high"? Our conceptual framework in Section 3 demonstrates this idea theoretically: high salary histories may signal that a worker's salary expectations may exceed that worker's value. Our survey results find supporting evidence for this: workers are less willing to disclose extremely high salaries.⁵¹ In later sections (7.4), we show this intuition to be correct: our recruiters denied callbacks to workers disclosing the highest salaries. Given this, silent candidates might be interpreted differently. Silent workers may contain a mixture of those whose salaries are too low as well as some whose salaries are too high. This may partly explain why our recruiters do not assume that silent candidates are 0th percentile workers; in principle, some of the silent types could be higher percentile workers hoping to avoid appearing overpriced.

We also find a second new mechanism behind the lack of full punishment: recruiters may believe that disclosure/silence choices are not driven entirely by a candidate's current salary level. Some workers may be inherently uncomfortable disclosing. They may feel that revealing salary history and/or allowing it to be used in hiring is repugnant. Our own survey— and others like it— suggests that women in particular do not like disclosing. Women are about 12 percentage points less likely to disclose than men when not prompted, and women disclose about 1 percentage point less often than men when prompted.⁵² Our survey results suggest that women are less willing to disclose their salaries, even after controlling for salary, education, and other characteristics. Recruiters may anticipate this and realize that female nondisclosure may be less informative. If women dislike disclosing—whatever their incomes are—then we should find that female nondisclosers would contain a greater variety of types (and not just low-salary types). To recruiters who understand this, the implication is that female nondisclosers may not signal a lower historical salary.

This is exactly what we find. In the even columns of Table 2, we interact disclosing with gender. Our results show that women job candidates are punished *less* for nondisclosure. Recruiters penalize a silent man's WTP by \$5,700 more, his outside option by \$9,200 more, and his offer by \$6,900 more than for a silent woman. Thus, recruiters change their beliefs less about a nondisclosing female compared to a nondisclosing male. The flip side of this behavior, however, is that the benefit of disclosing is also smaller for women.

From the perspective of our first mechanism, this result is particularly striking. Female candidates—

⁵¹See Agan, Cowgill and Gee (2020) and Cowgill et al. (2021).

⁵²The difference when prompted is not statistically significant.

who are paid less—should have a *lower* risk of being "too high" for an employer's budget. Silent women should therefore make up a greater proportion of "too low" candidates, and thus be penalized *more* for remaining silent in their beliefs about their value, outside options, and offers. However, we find the opposite. This is suggestive evidence that our recruiters are somewhat sophisticated consumers of disclosure information and are aware that women perceive a higher psychic cost of disclosing. Some 75% of our recruiters are themselves female (and 69.7% of recruiters in the economy broadly),⁵³ and 71% have three-plus years of prior recruiting experience (Table **??**). For these reasons, recruiters may understand that nondisclosing women may not be hiding bad information, but may simply be less willing to share their prior salary.

One potential alternative explanation for our results is that recruiters simply misjudged the average market wages for this job. We explore this in the appendix (Section F) and find that lack of knowledge does not seem to be driving our results. This is both because recruiters reacted to lack of information by doing more research about average market wages, and because our results are similar when we create a subset on recruiters who see half their job candidates disclosing.

7.2 Amounts Disclosed

Our second set of results is about the amounts disclosed among those who disclosed. In the odd columns of Table 4, we use Equation (6) to measure how recruiters reacted to an additional \$1 of salary a candidate disclosed.

On average, higher salaries increase recruiter WTP for candidates. This is consistent with employers believing that prior salary carries information about worker quality. However, WTP does not increase 1:1 with current salary—in column 1, we see that, on average, for every additional \$1 of current salary disclosed, WTP increases by \$0.65. Column 3 shows that recruiters' expectations about the candidates' outside offers similarly increase with higher amounts. The effects on competing offers is actually more responsive to disclosure than WTP. For every +\$1 of current salary disclosed, this median increases by \$0.77 (compared to \$0.65 for WTP, statistically significantly different with p = 0.002). In aggregate, salary offers have a stronger relationship with willingness to pay than with outside offers (Table G3). However, disclosures about salary history have a bigger effect on competing offers. This is consistent with the idea that salary history contains information about the worker's outside option, even if it is less informative about the worker's quality.

This idea—that salary history contains more information about worker quality than about outside offers—appeared earlier in our conceptual framework. In Section 3, we showed that if salary history was more informative about outside offers than about value, then high-salary-history workers would yield the lowest surplus and well-paid workers would face a trade-off between getting an

 $^{^{53}{\}rm See} \ {\rm https://www.bls.gov/ooh/business-and-financial/human-resources-special ists.htm.}$

offer and getting a high offer. Our results here foreshadow this result, and we will direct evidence in a later section.

In the even columns of Table 4, we also examine whether amounts disclosed have a different effect for male and female candidates. In aggregate, do recruiters respond to an extra dollar disclosed from men and women similarly? For most of the outcomes in Table 4, we do not find evidence that extra total dollars disclosed by female candidates are treated differently from those disclosed by male candidates. The coefficients on "Female x Disclosed x Amt Disclosed" are all statistically insignificant and small. Our results here are highly aggregated, and we decompose salary variation into multiple sources using Equation (6) next.

7.3 Heterogeneous Effects of an Extra +\$1

Variation in current salaries can arise from several sources. Some workers in our experiment work at firms with high or low overall wage distributions. Some workers are on the high or low ends of those internal distributions. Last, some are female and earn less as a result of the gender wage gap. We now examine whether these variations in salary amounts have identical effects. Table 5 presents the results from Equation (6), decomposing our $AmountDisclosed_{i,j}$ variable into multiple sources of salary variation.

We broadly find that not all dollars are created equally. As predicted in our theory, within-firm variation—labeled "+\$10K within Firm" in Table 5— is the most informative to recruiters in the sense that it has the steepest slope: each additional \$1 disclosed from this within-firm variation increases WTP by \$0.70. In theory, this variation is the only true source of surprise to the recruiter, as it is uncorrelated with anything else on the application form. By contrast, recruiters could anticipate that a worker at Apple (a high-wage firm in our sample) is paid around a certain average ("+\$10K from Firm"), and/or that men are typically paid more ("+\$10K from Male"). Although it is not always different from the other coefficients in a statistical sense, the slope on the firm's component is nonzero and is the second steepest across most of our regressions, suggesting that recruiters do not fully anticipate average firm wage differences.

The gender slope is the flattest and least informative: for each additional \$1 candidates disclose because of the gender wage gap which favors men, recruiters give offers that are \$0.48 higher. Our results show that compared to other sources of variation, recruiters discount extra dollars given to men. They may interpret that men are paid extra for spurious reasons and thus find their higher salaries less informative about quality or the candidate's outside offers. Alternatively, they may feel that men's higher reported salaries are uninformative because they are exaggerated lies.

In fact, the gender slope is particularly small (flat, uninformative) for our measure of perceived worker quality (WTP). For every +\$1 given for within-firm reasons, recruiters increase WTP by

\$0.70. However, recruiters only increase WTP by \$0.42 for every \$1 allocated to men through the gender wage gap (WTP vs. outside option p = 0.00). Despite lower average amounts, the gender slopes—including those on candidate quality (WTP)—are still far from zero. Even if recruiters update less from the gender component, they still impact the callback and compensation decisions in our experiment.

The slope for an additional \$1 from being male is steeper on our measure of perceived outside options, although still less than the other slopes (and still statistically different from zero). While recruiters are reluctant to believe men's higher salaries are a signal of quality, they are more willing to believe they are correlated with the candidate's outside offer.⁵⁴ For all three components of salary variation, the slope on the recruiter's WTP is lower than the slope on beliefs about the candidate's outside options.

7.4 Decision-Making Outcomes: Whom to Hire and How to Pay

Until now, our empirical results have been about employer beliefs and whether recruiters have more favorable (or unfavorable) beliefs about candidates depending on their characteristics and choices. Our final set of results is about how recruiters synthesize inferences into decisions about whom to move forward with hiring and how to compensate them. These decisions are modeled as a function of beliefs in Equations 1 (salary amounts) and 2 (callbacks). As we covered above, the effects on salary offers are straightforward: disclosing workers—particularly those with high salaries—receive higher salary offers.

Our results about callbacks are more nuanced. On average, our results about callbacks generally go in the *opposite* direction as those on salary offers. Disclosing workers—especially men and workers with high salaries—are less likely to be recommended *at all*. Although they enjoy higher salary offers when selected, they are less likely to move forward in the hiring process.

Our estimates in this section are sometimes underpowered because of the binary nature of the outcome variable. In addition, the relationship between callbacks and amounts disclosed is noisy near the lower end of the salary distribution. Where possible, we use the theoretically motivated outcomes we collected—employer surplus, probability of acceptance, and others—to help shed additional light on the story.

⁵⁴Higher outside offers could come partly from a candidate's search effort, or from rival companies' search efforts targeting the candidate. When we try to measure the direction of search in Table G8, we find mostly small and insignificant slopes. However, the direction suggests that recruiters believe men have higher outside options through higher candidate-driven search. By contrast, the other slopes are also correlated with greater outside offers. Nevertheless, recruiters believe these offers originate from rival companies' search efforts, rather than from the candidates'.

7.4.1 Callbacks and Amounts

Our strongest evidence about disclosures and callbacks comes from our specifications that include amounts. Across several specifications, disclosing a higher salary reduces callbacks. In Figure 2, we visualize the relationship between amounts disclosed and callbacks using our full data set. As the figure reveals, we find a noisy relationship in the lower $\approx 15\%$ of the data (below \$70K). However, in the upper $\approx 85\%$ of our data, there is a visibly negative relationship.⁵⁵

Tables 9 and 10 present these results as regressions, using Equation 6 for callbacks and outcomes theoretically related to callbacks (employer surplus, predicted probability that candidate accepts offer, and probability of acceptance \times surplus). Table 9 includes our full sample. Even in our full sample, which includes the noisy results on the low salaries, the relationship is downward sloping (albeit small and statistically insignificant). When we focus on the upper 85% of the data in Table 10, we find the stronger negative relationships visible in Figure 3. Every +\$10,000 extra disclosed by a candidate reduces the probability of a callback by 4 percentage points (6.25% over the mean for nondisclosing candidates).

Table 9 shows that higher salaries not only reduce callbacks but also the key drivers of the callback decision. Every \$1 extra disclosed by a candidate reduces employer surplus by \$0.03. Higher disclosure amounts also reduce employer surplus, as well as beliefs about the candidate accepting the offer. Table 10 also presents gender interactions for our slopes. We find that the amount/callback relationships are generally stronger for men, both statistically and economically. This is in contrast to the results in our earlier section, in which female amounts featured a higher slope.

As we showed above, we can disaggregate the effects of salary differences. Table 8 examines callback-related outcomes using Equation 6 (using the full sample). This allows us to measure the effects of a +\$1 on callback outcomes, depending on whether it came from the gender wage gap, between-firm variation, or within-firm variation. For nearly all of our callback-related outcomes and sources of variation in Table 8, we find negative relationships with extra dollars. Randomly assigned higher amounts from any source we study are correlated with lower callback outcomes, although many of these effects are statistically insignificant.

Of these, our results on the gender wage gap (+\$10K from being male) are especially precisely

⁵⁵To find the 70K cutoff, we ran regressions predicting callback as a function of salary disclosed and the square of salary disclosed from versions of Equation 5 systematically creating subsets for the data from those above 60K, 70K, 80K, and so on. The squared term is insignificant after we subset to 70K and above. We also ran a regression predicting callback as a function of salary disclosed from versions of Equation 5, systematically subsetting the data from those above 60K, 70K, 80K, etc. We found the coefficient on the amount disclosed is negative and statistically significant for amounts \$70K and above. Results are available from the authors upon request. This threshold may have arisen at \$70K because of the following statement in our instructions: "Software engineers currently at our firm make between \$70,000 and \$120,000. You should not feel constrained by our current range, and we welcome your own research about what candidates should be paid."

estimated. We find that +\$10,000 extra salary given to men because of the gender wage gap causes a 5 percentage point *decrease* in the probability of a callback and destroys \$7,000 in employer surplus. The probability of accepting the TIOLI offer (and its product) also declines. These results show that the higher wages afforded to men (through the gender wage gap) harm their callback chances when disclosed. Men's higher salaries appear to lower their salary-net attractiveness as candidates, in some cases by more than other sources of variation. This again suggests some level of sophistication by recruiters.

This is not a contradiction of our previous results. To the contrary, it is the natural extension of the findings established earlier and in our conceptual model. As our survey and theoretical results show, callback decisions are not entirely a function of worker quality; they must also incorporate expected costs. As discussed earlier in our paper, salary disclosures may also affect whether a candidate is hired at all (and by whom), versus leaving the position unfilled or hiring a different candidate.

Throughout several tests above, we find that salary history disclosure increases beliefs about outside offers more than candidate quality (Tables 2, 3 and 5). Where could such a process lead? As outside offers outpace candidate quality, eventually employers' margins will be squeezed. As salary expectations rise too high, selecting an expensive candidate becomes a bad economic deal. Employers must pay a premium to keep the candidate out of a rival's hands, but the candidate's (perceived) quality may not warrant it. Our findings about callbacks are logical extensions of our earlier findings.

7.4.2 Silence and the Choice to Disclose

We now turn to how silence affects callbacks. In typical models of disclosure, the negative interpretation of silence is theoretically related to higher disclosure amounts signaling better news. We see this pattern in our results in Sections 7.2 and 7.3: Higher disclosure amounts contain good information, and lower disclosures (and thus silence) are interpreted negatively. For callbacks, higher disclosure values appear to signal *worse* news to employers. This raises the possibility that silence would actually have a *positive* effect for callback decisions. We examine this in Table 6 using our full sample, and Table 7 using our sample that disclosed above \$70K.

Our results on this question are imprecise but lean closer to the conclusion that "silence is *good* for callbacks" than the opposite, particularly for men and highly paid candidates. Although our results about the callback variable itself are never statistically significant, we do sometimes find statistically significant results about employer surplus, and surplus $\times p(\text{accept})$. Although silent workers are regarded as lower quality (Table 2, column 1, discussed above), they are also cheaper to keep away from competing firms and opportunities (Table 2, column 3, also discussed above).

When we examine these results heterogeneously by gender, we find that female silence has a

smaller effect on callbacks than male silence. Stated oppositely, on average, women's callbacks are punished less from disclosures than male callbacks. In fact, our gender interactions contain no statistically significant evidence of women being punished at all (on average) for disclosing. This is consistent with our earlier finding that a woman's disclosure/silence choices contain less information about her underlying value as an employee.

8 Effects of Salary History Bans

Until now, our paper has addressed how employers react to salary history disclosures. We now examine what our findings suggest about the public policies motivating our study: salary history bans. The goal of these bans is to equalize outcomes by suppressing disclosures. However, bans cannot completely suppress disclosures. They forbid employer prompting, but not *voluntary* disclosure. Analyzing the effects of bans requires data and assumptions about worker compliance.

In this section, we combine the results of our field experiment with a survey of the working public in the United States. Our survey included approximately 1,000 individuals representative of the working American public (defined as Americans in the labor force between the ages of 22 and 55) and helps us identify which candidates are more likely to disclose unprompted (or to refuse disclosure, even when asked). Here, we combine the main results of the survey with our field experiment with recruiters to study implications for salary history ban policies.

As described in our theory section (3.5), we model the ban as an increase in the cost of disclosure. As such, the prompt affects who discloses and how disclosures are interpreted. These disclosures then affect employers' choices through the mechanisms we document above. Because we have already explored these mechanisms empirically and theoretically, this section focuses on the bottom-line effects of the ban. Our design allows us to separate the two major *designs* of the bans: the "full ban" (no salary questions, ever) and the "partial ban" (which allows salary history questions after the initial offer).

8.1 Estimation Strategy

Our strategy for simulating the potential impacts of a salary history ban contains several components.

8.1.1 Differentiating Prompted vs. Unprompted Disclosures

We begin by measuring whether recruiters interpret salary disclosures differently, depending on whether job candidates *volunteered*, rather than provided disclosure in response to a question. We analyze this question in depth in Appendix E.1 and summarize here: The differences between

prompted and unprompted disclosures are relatively small and cannot be rejected from zero. We *can* reject large effects. If anything, silence in response to a prompt lowers beliefs about candidate quality (compared to silence without a prompt). This small difference does not mean that prompts do not matter. Rather, it suggests that the prompts affect outcomes mainly through the effects on worker disclosure behavior, not through the interpretation of disclosures.

8.1.2 Outcomes

One could combine outcomes in a number of ways to represent workers' overall well-being. Instead, we separately present results about callbacks and salary amounts (conditional on a callback).⁵⁶ By presenting results on callbacks and salary amounts separately, we invite readers to import their own policy objectives about how these outcomes should be weighed. Even callbacks attached to low salaries may still be useful to workers if they grant flexibility and/or negotiating power with a current employer.⁵⁷

8.1.3 Differentiating Full vs Partial Ban

As discussed previously, there are two types of bans, which we call "partial" and "full" bans. Partial bans prevent employers from seeking histories only until the first offer has been made. In contrast, full bans prevent salary histories from *ever* being sought. Under both ban scenarios, candidates can volunteer their salary information at any time.

For the outcomes about who is called back, outcomes for a full or partial ban are the same. Callback decisions under the full or partial bans are made only using unprompted information. However, choices about salary can differ between the full and partial ban. In the partial ban, employers can ask salary-history questions after making the first offer. As such, we modify our original data to simulate the effect of a partial ban. We use the callback decisions from our observations where there is no prompt on the application, and we use the offers from the observations where salary histories were asked for with a prompt. This simulates the scenario where an employer cannot ask for salary history information before the choice of whom to call back is made. But the employer *can* ask afterward about salary history and utilize any resulting information before making a final offer.

⁵⁶This means we restrict our sample to those who would be suggested for a callback and look at the offers they might enjoy were they to be made a job offer.

⁵⁷One could instead present callbacks \times offer amount ("expected salary"). However, the outside option for a searching worker might not be zero if a job application fails. Many workers, including our fictitious candidates, might be searching for a job while currently employed, and, as such, lack of a callback is not a zero outcome.

8.1.4 Regression specifications

To estimate the effect of asking salary history questions with a prompt, we use Equation 9, below. Standard errors are robust and clustered at the recruiter level. The equation contains interactions that estimate the effects of these questions separately for men and women.

$$y_{i,j} = \alpha + \beta_1 SalaryHistoryAsked_j + \beta_2 Female_{i,j} + \beta_3 Female_i \times SalaryHistoryAsked_j + \epsilon_j$$
(9)

Notice that Equation 9 does not include a variable for whether the candidate disclosed. This is because our strategy in this section is to model candidates' disclosure decisions—and employers' inferences from them—as potentially downstream from the prompt. As described above, simulations of full or partial bans use the same observations if the outcome, $y_{i,j}$, is a callback, but different sets of observations when the outcome is later in the hiring process (like a TIOLI amount for those called back).

8.1.5 Regression weights

We estimate Equation 9 using weights to regulate levels of worker compliance. Because our candidates are fictitious, we do not know how they would respond to a salary history prompt (or its removal) in real life. However, our experiment measures recruiters' reactions to all sets of compliance behaviors for each candidate, both with and without the prompt.

We can then examine how the effect of the ban changes with assumptions about compliance. To show an example of this use of weights, suppose we wanted to assess the ban assuming that all subjects were "compliers." We would assign a weight = 1 for all observations in which the candidate is prompted and discloses, and (similarly) a weight = 1 for observations in which the candidate is *not* prompted and does not disclose. All other observations would receive a weight of zero.⁵⁸ In this example of compliers, the coefficient on the prompt captures the effect of all workers changing from nondisclosers to disclosures when prompted. In footnote 59, we review a weighting scheme as if everyone were an "always-discloser."⁵⁹

We aim to employ weights for Equation 9 reflecting the true probabilities of the disclosure behavior for each of our candidates. If women like Jessica from Oracle (Candidate ID No. 5 in our experiment) tend to be compliers, we would place a greater weight on her observations that include disclosure (when prompted) and nondisclosure (when not prompted). Below, we discuss

 $^{^{58}}$ This effectively drops all observations in which a candidate is asked and stays silent, or volunteers without asking. 59 Suppose we wanted to assess the ban assuming that all subjects were "always-disclosers." We would receive a weight = 1 for all observations in which the candidate is prompted and discloses, and, similarly, a weight = 1 for observations in which the candidate is *not* prompted and yet discloses (volunteers unprompted). All other observations would receive a weight of zero. This effectively drops all observations in which a candidate does not disclose, leaving only observations where the candidate discloses (either with prompting, or without).

our strategy for obtaining the weights. Using these weights, we can use Equation 9 to estimate how the ban will affect aggregate outcomes of our experiment (for example, the overall level of men's and women's salary offers, and their ratio).

8.2 Surveys of Workers and Regression Weights

To estimate each of our candidates' likely compliance behavior, we conducted a survey of over 1,000 American workers ages 22 to 55 using a survey company.⁶⁰ The surveys asked respondents about their demographics and whether they would disclose their salary when asked (or volunteer when not asked).⁶¹ Using this data, we can identify each survey respondent's disclosure type ("always-discloser," "never-discloser," etc.) by asking about candidate disclosure behavior when prompted by an employer (and when not). We also asked how workers in the survey volunteer (or not) in scenarios where asking is illegal.⁶² Our survey also asked subjects to identify their gender, their overall income, their industry and occupation, and whether they were relatively well paid (or not) compared to other people in the same job at the same company.

These covariates helped us link survey respondents to fictitious candidates in our experiment with similar characteristics. Using these data points, we developed a map between our survey responses and the characteristics of our job candidates. Table G12 displays some descriptive regressions of these mappings. Always-disclosers make up 28% of our survey sample. They are more likely to be male and are slightly higher-paid.⁶³ 52% of our survey sample are compliers. Compliers have the opposite set of correlations: more female and in lower-paying occupations, lower-paid conditional on jobs and industries, and less likely to be paid more than peers. Never-disclosers make up 19% of our sample; we find they are more likely to report high salaries within

⁶⁰We used the company Prolific Academic, https://www.prolific.co/. In a short technical report (Agan, Cowgill and Gee, 2020) and a longer paper (Cowgill et al., 2021), we document this survey instrument and analyze the data included in this paper in more detail.

⁶¹The exact wording of the latter question was, "Imagine that no one involved in the hiring process has asked you about your most recent salary. However, you can legally disclose this information voluntarily. Would you tell them your most recent salary?" For the first question, it was, "Imagine it is perfectly legal for someone involved in the hiring process to ask your most recent salary. If someone asks, would you tell them your most recent salary?" For a subset of respondents, we have information about their choices in two real job searches in which they were asked and not asked for salary history. In addition to these hypothetical scenarios, we also asked about the subjects' actual disclosures when they encountered questions in their job searches. Answers to the hypothetical questions and real questions are positively correlated (results available from authors upon request). We also found that subjects' choices to disclose were approximately the same if employers were banned from asking vs. if employers were allowed to ask but did not.

 $^{^{62}}$ When salary questions have been banned, workers may be more (or less) willing to volunteer—compared to settings where they are allowed but companies decline to ask. Our surveys explicitly address this possibility. We find that when an employer does not ask, about 70–74% of workers do not disclose, regardless of three scenarios: 1) whether asking is banned by law, 2) if it is legal but the employer chooses not to, or 3) if the legality is ambiguous (Cowgill et al., 2021).

⁶³They work in slightly higher-paying jobs and industries.

their firm but to work in lower-paying occupations.

Using this data, we construct eight cells ({male, female} \times {High-Wage Firm, Low} \times {High Salary within Firm, Low}). The distribution of disclosure types of all eight groups can be browsed in Table 11. The proportion of the population in each of these cells is relatively uniform, between 11.1% and 15.4%. Across these cells, women are more likely to be compliers, and are less likely to be always-disclosers. Respondents working in high-wage firms are more likely to be always disclosers and less likely to be compliers. Within each cell, never disclosers are always the smallest proportion of a compliance type. Each of the candidates in our field experiment can be mapped back to one of these cells. We thus merge the contents of Table 11 with our experimental data, giving weights to the experimental observations that the survey indicates are more likely. The final columns of Table 11 summarize how much each group gains/loses from disclosing (on average).

8.3 Ban Simulation Results

8.3.1 Ban Effects on Callbacks

Table 12 shows the effects of banning employers from prompting salary-history disclosure on whether women and men in our field experiment are selected to receive a callback. Under either a partial or a full ban, the recruiter would only observe the salary information on the job application without a prompt at the time that callback choices are made, and as such, the results are the same under either type of ban. Either with or without a ban, 64% of women and 63% of men are recommended for a callback. There is no gender gap in callbacks before the ban, and there continues to be no gender gap after the ban in our setting with a preexisting gender wage gap.

These results are similar to a number of studies which observe small or nonexistent changes in employment or job changing from bans (Bessen, Meng and Denk, 2020; Sinha, 2019; Hansen and McNichols, 2020; Mask, 2020).⁶⁴

8.3.2 Effects of a Ban on Salary Offers Conditional on Callback

Table 13 presents the effects of full bans, where salary disclosure may never be prompted, and partial bans, where an inquiry may be made later in the hiring process. The story is similar for both types of bans: Bans close the gender gap. In the left-hand panel of Table 13, we see that the ratio of annual salary conditional on callback for women to men is 0.91 before a ban, and then rises to 0.97 after a ban (p = 0.00), meaning that women and men are almost equal in annual salary after the ban is in place. A partial ban has the same pattern of effects, but the magnitude is significantly

⁶⁴Some other studies find changes in job transitions, either overall (Sran, Vetter and Walsh, 2020) or for those who entered the labor force during a recession (Mask, 2020).

smaller and less precise. The ratio of female to male is 0.91 before a ban and rises to 0.92 after a ban (p = 0.13).⁶⁵

However, how bans close the gender gap is important. Bans lower salaries for both women and men. But a full ban harms men by an average of \$8,299 (p = 0.00) while women only lose \$1,447 (p = 0.45). The results are even more harmful to women for a partial ban, with women losing \$3,858 (p = 0.00) and men losing \$6,159 (p = 0.00) on average. This same pattern can be seen for alternative outcomes like salary×callback (Table G13). For either type of ban, the gender gap is closed by greater harm to men than to women. Policymakers may have hoped that salary history bans would raise the salary of women, but our field experiment shows women are worse off and men are much worse off.

In our section on effects of disclosure (Section 7), we found that silence lowered employers' beliefs about candidate quality and (especially) competing offers. This resulted in lower salary offers. It is possible that a salary history ban would have no effect on the levels of silence vs. disclosing (i.e., if the world was made up of only always-disclosers and never-disclosers). However, our survey shows that about 52% of the U.S. workforce are ban compliers, and as such, bans increase silence. Bans lower inequality but do so by increasing silence. The resulting silence harms men more than women. In short, bans divide the pie more equally between male and female job candidates, but they also shrink the total size of the pie.

9 Conclusion

This paper assesses salary disclosure microeconomically. We develop a conceptual framework that adapts the voluntary-disclosure literature to hiring and wage-setting in labor markets. We then deploy a novel field-experimental design—a *two-sided audit*—that utilizes recruiters to peer inside the black box of hiring. Our design permits us to vary characteristics of job applicants and of firms. Our experiment traces disclosures from messages to updated beliefs, and from employer inferences to choices that affect candidates' pocketbooks and daily lives. The tools we develop could be used to examine a wide variety of other interventions that alter firm hiring practices.

We specifically look at the effects of salary disclosure on the hiring process. Salary disclosures are of interest because salary history bans have become a popular policy instrument to close gender (and other) wage gaps. Our results would be difficult to obtain from observational data sets, where one does not usually observe whether a job candidate has disclosed his or her salary, and where the decision to disclose is endogenous and likely correlated with possible confounders like current salary and gender. Using a field experiment allows us an unprecedented level of observation and

⁶⁵The ratio of female to male salaries disclosed is 0.85 in our experiment by construction. Under all the policy regimes studied, the recruiters in our experiment narrowed the wage gap.

control over correlates of disclosure.

We have three main results. First, silent job candidates are believed to be less valuable and to have lower outside offers, and so are extended lower take-it-or-leave it offers. Second, disclosing is especially beneficial to those with higher salaries. In settings with a gender wage gap, men benefit more from disclosing than women, in part through their higher salaries. Additionally, women feel more discomfort from disclosing, and knowing this, employers may update less strongly in response to silence from a woman than a man. Both the gender wage gap and other nonwage factors mean the benefits of disclosing accrue mostly to men. Third, unlike other settings where more is always better, at some point a salary disclosure can be so high that it makes a worker unattractive to the firm. This leads to our third result, which is that for most people, disclosing a higher amount lowers the likelihood of obtaining a callback. Our results highlight the trade-offs of disclosing high numbers, which results in higher offers but a lower likelihood of making it to the next step in the hiring process.

Next, we pair our field experiment with a survey of U.S. workers to simulate the effects of salary history ban policies on gender gaps. Our survey allows us to estimate how workers will comply, or not, with salary history ban attempts to lower the number of salary disclosures. In our setting with a preexisting gender wage gap, we find that salary history bans have little effect on callback rates for men or women. In our first set of results, we find that staying silent was generally harmful for outcomes beyond callbacks. Bans are meant to increase silence about salary history, and so it is no surprise that bans, which prohibit ever asking for salary history, decrease salary offers for men and women. However, these bans also decrease the gender gap, because the harm of silence is greater to men than to women. In our setting, we find that salary history bans obtain their stated goal of increasing equity, albeit through harming female workers but harming male workers more.

Our paper leaves key questions for future research: First, we do not study whether employers are "correct" in the inferences we document. Are low-disclosing workers truly worse performers? Are they silent about their salaries for strategic reasons? Do they truly have poorer outside offers? The subjects in our field experiment are experienced recruiters who have incentives to make accurate choices. However, we cannot independently assess their judgments. Any such assessment would be challenging: it would require worker-level performance data, including for rejected candidates (to assess whether they had been rejected mistakenly). For these reasons, the beliefs we document may persist, even if they are incorrect.

In addition, it would require knowledge of employers' subjective map between worker-level behaviors and firm utility. These measurement challenges prevent *actual* employers from assessing their own workers; prior research suggests that employer learning can be slow in practice (Lange, 2007; Kahn and Lange, 2014). According to a 2019 survey by the Society of Human Resources Management, 77% of employers do not measure the quality of hires on a regular basis.⁶⁶ Employers may lack corrective feedback, may be inattentive or otherwise fail to learn for behavioral reasons (Mullainathan, Schwartzstein and Shleifer, 2008; Jin, Luca and Martin, 2015).

Second, we do not answer whether workers are holistically better off, or more equal, as a result of the interventions we study. Our results speak to employer demand: whether workers are paid more, given higher salary offers, or given an offer at all. Combining these into a single measure of worker welfare involves trade-offs. Several of our results go in opposite directions, for theoretically grounded reasons. Revealing a low salary can make a good worker appear cheap and make a callback more likely. Having any form of outside offer gives workers options and negotiating power. However, appearing cheap may also lower the compensation amounts attached to selections. Different workers and policymakers may prefer to resolve these trade-offs differently. Our paper offers information for understanding these trade-offs.

⁶⁶https://www.shrm.org/ResourcesAndTools/business-solutions/Documents/Talent -Acquisition-Report-All-Industries-All-FTEs.pdf

References

- Acemoglu, Daron. 1999. "Changes in unemployment and wage inequality: An alternative theory and some evidence." *American economic review*, 89(5): 1259–1278.
- Acemoglu, Daron. 2002. "Technical change, inequality, and the labor market." *Journal of economic literature*, 40(1): 7–72.
- Acemoglu, Daron, and David Autor. 2011. "Skills, tasks and technologies: Implications for employment and earnings." In *Handbook of labor economics*. Vol. 4, 1043–1171. Elsevier.
- Acquisti, Alessandro, and Christina M Fong. 2015. "An experiment in hiring discrimination via online social networks."
- Adler, Laura. 2020. "What's a Job Candidate Worth? Explaining Variation in Pay-Setting Practices." *Working Paper*.
- Agan, Amanda, and Sonja Starr. 2018. "Ban the Box, Criminal Records, and Racial Discrimination: A Field Experiment." *The Quarterly Journal of Economics*, 133(1): 191–235.
- Agan, Amanda, Bo Cowgill, and Laura Gee. 2020. "Do Workers Comply with Salary History Bans? Voluntary Disclosure, Adverse Selection, and Unraveling." Vol. 108, 33–37.
- Allon, Gad, Achal Bassamboo, and Ramandeep S Randhawa. 2012. "Price as a Signal of Product Availability: Is it Cheap?" *Available at SSRN 3393502*.
- Analysts, Staffing Industry. 2017. "RPO Market Developments."
- Asch, Beth J. 1990. "Do incentives matter? The case of navy recruiters." ILR Review, 43(3): 89–S.
- Ashenfelter, Orley C, Henry Farber, and Michael R Ransom. 2010. "Labor market monopsony." *Journal of Labor Economics*, 28(2): 203–210.
- Åslund, Olof, and Oskar Nordströum Skans. 2012. "Do anonymous job application procedures level the playing field?" *ILR Review*, 65(1): 82–107.
- **Autor, David H.** 2001. "Why do temporary help firms provide free general skills training?" *The Quarterly Journal of Economics*, 116(4): 1409–1448.
- Avivi, Hadar, Patrick M Kline, Evan Rose, and Christopher R Walters. 2021. "Adaptive Correspondence Experiments." National Bureau of Economic Research.
- **Barach, Moshe A, and John J Horton.** 2021. "How do employers use compensation history? Evidence from a field experiment." *Journal of Labor Economics*, 39(1): 193–218.
- Bartos, Vojtech, Michael Bauer, Julie Chytilova, and Filip Matejka. 2016. "Attention Discrimination: Theory and Field Experiments with Monitoring Information Acquisition." *American Economic Review*, 106(6): 1437–1475.
- Becker, Gary S. 1957. The economics of discrimination. University of Chicago press.
- **Becker, Gordon M, Morris H DeGroot, and Jacob Marschak.** 1964. "Measuring utility by a single-response sequential method." *Behavioral science*, 9(3): 226–232.
- Behaghel, Luc, Bruno Crépon, and Thomas Le Barbanchon. 2015. "Unintended effects of anonymous resumes." *American Economic Journal: Applied Economics*, 7(3): 1–27.
- Bertrand, Marianne, and Esther Duflo. 2017. "Field experiments on discrimination." In *Handbook of economic field experiments*. Vol. 1, 309–393. Elsevier.
- **Bessen, James E, Chen Meng, and Erich Denk.** 2020. "Perpetuating Inequality: What Salary History Bans Reveal About Wages." *Working Paper*.
- **Biasi, Barbara, and Heather Sarsons.** 2020. "Flexible wages, bargaining, and the gender gap." National Bureau of Economic Research.
- Blackaby, David, Alison L Booth, and Jeff Frank. 2005. "Outside offers and the gender

pay gap: Empirical evidence from the UK academic labour market." *The Economic Journal*, 115(501): F81–F107.

- Black, Ines, Sharique Hasan, and Rembrand Koning. 2020. "Hunting for talent: Firm-driven labor market search in the United States." *Working paper*.
- Blau, Francine D, and Lawrence M Kahn. 2006. "The US gender pay gap in the 1990s: Slowing convergence." *ILR Review*, 60(1): 45–66.
- **Blau, Francine D, Peter Brummund, and Albert Yung-Hsu Liu.** 2013. "Trends in occupational segregation by gender 1970–2009: Adjusting for the impact of changes in the occupational coding system." *Demography*, 50(2): 471–492.
- Bordalo, Pedro, Katherine Coffman, Nicola Gennaioli, and Andrei Shleifer. 2019. "Beliefs about gender." *American Economic Review*, 109(3): 739–73.
- Brands, Raina A, and Isabel Fernandez-Mateo. 2017. "Leaning out: How negative recruitment experiences shape women's decisions to compete for executive roles." *Administrative Science Quarterly*, 62(3): 405–442.
- Cahuc, Pierre, Fabien Postel-Vinay, and Jean-Marc Robin. 2006. "Wage bargaining with onthe-job search: Theory and evidence." *Econometrica*, 74(2): 323–364.
- **Cappelli, Peter, Steffanie L Wilk, et al.** 1997. *Understanding selection processes: organization determinants and performance outcomes.* Center for Economic Studies, US Department of Commerce, Bureau of the Census.
- **Clifford, Robert, and Daniel Shoag.** 2016. "'No More Credit Score': Employer Credit Check Bans and Signal Substitution."
- **Condren, Conal.** 1997. Encouraging Recruiter Achievement: A Recent History of Military Recruiter. Rand Corporation.
- Cowgill, Bo, Amanda Agan, , and Laura Gee. 2021. "Asking and Not Asking about Salary History: How do Job-Seekers React?" *Working Paper*.
- **Cowgill, Bo, and Patryk Perkowski.** 2021. "Delegation in Hiring: Evidence from a Two-Sided Audit." *Working Paper*.
- Croson, Rachel, and Uri Gneezy. 2009. "Gender differences in preferences." *Journal of Economic literature*, 47(2): 448–74.
- Cullen, Zoë B, and Bobak Pakzad-Hurson. 2021. "Equilibrium Effects of Pay Transparency." Working Paper.
- Davis, Jesse, Paige Ouimet, and Xinxin Wang. 2020. "Do Salary History Bans Limit Discrimination?" *Working Paper*.
- **Dee, Thomas S.** 2005. "A teacher like me: Does race, ethnicity, or gender matter?" *American Economic Review*, 95(2): 158–165.
- **Dey, Matthew S, and Christopher J Flinn.** 2005. "An equilibrium model of health insurance provision and wage determination." *Econometrica*, 73(2): 571–627.
- **Doleac, Jennifer L., and Benjamin Hansen.** 2020. "The Unintended Consequences of "Ban the Box": Statistical Discrimination and Employment Outcomes When Criminal Histories Are Hidden." *Journal of Labor Economics*, 38(2): 321–374.
- **Dranove, David, and Ginger Zhe Jin.** 2010. "Quality disclosure and certification: Theory and practice." *Journal of Economic Literature*, 48(4): 935–63.
- Eckel, Catherine C, and Philip J Grossman. 2001. "Chivalry and solidarity in ultimatum games." *Economic Inquiry*, 39(2): 171–188.
- Exley, Christine L, and Judd B Kessler. 2019. "The gender gap in self-promotion." National

Bureau of Economic Research.

- Exley, Christine L, Muriel Niederle, and Lise Vesterlund. 2020. "Knowing when to ask: The cost of leaning in." *Journal of Political Economy*, 128(3): 816–854.
- Faberman, R Jason, Andreas I Mueller, Ayşegül Şahin, and Giorgio Topa. 2017. "Job search behavior among the employed and non-employed." National Bureau of Economic Research.
- Flory, Jeffrey A, Andreas Leibbrandt, and John A List. 2015. "Do competitive workplaces deter female workers? A large-scale natural field experiment on job entry decisions." *The Review of Economic Studies*, 82(1): 122–155.
- Fryer Jr, Roland G. 2007. "Belief flipping in a dynamic model of statistical discrimination." *Journal of Public Economics*, 91(5-6): 1151–1166.
- Gee, Laura K. 2019. "The more you know: Information effects on job application rates in a large field experiment." *Management Science*, 65(5): 2077–2094.
- **Gittleman, Maury, and Brooks Pierce.** 2015. "Pay for performance and compensation inequality: Evidence from the ECEC." *ILR Review*, 68(1): 28–52.
- Giuliano, Laura, David I Levine, and Jonathan Leonard. 2009. "Manager race and the race of new hires." *Journal of Labor Economics*, 27(4): 589–631.
- Glassdoor. 2019. "Progress on the Gender Pay Gap: 2019." White Paper.
- Goldfarb, Avi, and Catherine Tucker. 2012. "Shifts in privacy concerns." *American Economic Review*, 102(3): 349–53.
- Goldin, Claudia. 1990. "The gender gap: An economic history of American women." *Cambridge University Press, New York, Estados Unidos.*
- **Goldin, Claudia, Sari Pekkala Kerr, Claudia Olivetti, and Erling Barth.** 2017. "The expanding gender earnings gap: Evidence from the LEHD-2000 Census." *American Economic Review*, 107(5): 110–14.
- **Grossman, Sanford J.** 1981. "The informational role of warranties and private disclosure about product quality." *The Journal of Law and Economics*, 24(3): 461–483.
- Grossman, Sanford J, and Oliver D Hart. 1980. "Disclosure laws and takeover bids." *The Journal of Finance*, 35(2): 323–334.
- Hall, Robert E, and Alan B Krueger. 2012. "Evidence on the incidence of wage posting, wage bargaining, and on-the-job search." *American Economic Journal: Macroeconomics*, 4(4): 56–67.
- Hansen, Benjamin, and Drew McNichols. 2020. "Information and the Persistence of the Gender Wage Gap: Early Evidence from California's Salary History Ban." *National Bureau of Economic Research Working Paper*.
- Jin, Ginger Zhe, Michael Luca, and Daniel Martin. 2015. "Is no news (perceived as) bad news? An experimental investigation of information disclosure."
- Jovanovic, Boyan. 1979*a*. "Firm-specific capital and turnover." *Journal of political economy*, 87(6): 1246–1260.
- Jovanovic, Boyan. 1979b. "Job matching and the theory of turnover." *Journal of political economy*, 87(5, Part 1): 972–990.
- Kahn, Lisa B. 2010. "The long-term labor market consequences of graduating from college in a bad economy." *Labour Economics*, 17(2): 303–316.
- Kahn, Lisa B, and Fabian Lange. 2014. "Employer learning, productivity, and the earnings distribution: Evidence from performance measures." *The Review of Economic Studies*, 81(4): 1575–1613.

- Kessler, Judd B., Corinne Low, and Colin D. Sullivan. 2018. "An Audit Alternative: Measuring Employer Preferences and Beliefs without Deception." *Working Paper*.
- Khanna, Shantanu. 2020. "Salary History Bans and Wage Bargaining: Experimental Evidence." *Labour Economics*, 101853.
- Kline, Patrick, and Christopher Walters. 2019. "Audits as Evidence: Experiments, Ensembles, and Enforcement." *arXiv preprint arXiv:1907.06622*.
- Krueger, Alan B. 1993. "How computers have changed the wage structure: evidence from microdata, 1984–1989." *The Quarterly Journal of Economics*, 108(1): 33–60.
- Lange, Fabian. 2007. "The speed of employer learning." *Journal of Labor Economics*, 25(1): 1–35.
- Laschever, Sara, and Linda Babcock. 2003. Women Don't Ask: Negotiation and the Gender Divide. Princeton University Press.
- Lazear, Edward P. 2009. "Firm-specific human capital: A skill-weights approach." Journal of political economy, 117(5): 914–940.
- Levy, Frank, and Richard J Murnane. 1996. "With what skills are computers a complement?" *The American Economic Review*, 86(2): 258–262.
- Makridis, Christos, and Maury Gittleman. 2020. "On the Cyclicality of Real Wages and Employment: New Evidence and Stylized Facts from Performance Pay and Fixed Wage Jobs." *Available at SSRN 3017034*.
- Manning, Alan. 2003. *Monopsony in motion: Imperfect competition in labor markets*. Princeton University Press.
- Marianne, Bertrand. 2011. "New perspectives on gender." In *Handbook of labor economics*. Vol. 4, 1543–1590. Elsevier.
- Mask, Joshua. 2020. "Salary History Bans and Healing Scars from Past Recessions."
- Mason, Malia F, Alice J Lee, Elizabeth A Wiley, and Daniel R Ames. 2013. "Precise offers are potent anchors: Conciliatory counteroffers and attributions of knowledge in negotiations." *Journal of Experimental Social Psychology*, 49(4): 759–763.
- Mathios, Alan D. 2000. "The impact of mandatory disclosure laws on product choices: An analysis of the salad dressing market." *The Journal of Law and Economics*, 43(2): 651–678.
- Mazei, Jens, Joachim Hüffmeier, Philipp Alexander Freund, Alice F Stuhlmacher, Lena Bilke, and Guido Hertel. 2015. "A meta-analysis on gender differences in negotiation outcomes and their moderators." *Psychological bulletin*, 141(1): 85.
- Meli, Jeffrey, and James C Spindler. 2019. "Salary History Bans and Gender Discrimination." *Available at SSRN 3361431*.
- Milgrom, Paul, and John Roberts. 1986. "Price and advertising signals of product quality." *Journal of political economy*, 94(4): 796–821.
- Milgrom, Paul R. 1981. "Good news and bad news: Representation theorems and applications." *The Bell Journal of Economics*, 380–391.
- Moss-Racusin, Corinne A, John F Dovidio, Victoria L Brescoll, Mark J Graham, and Jo Handelsman. 2012. "Science faculty's subtle gender biases favor male students." *Proceedings of the National Academy of Sciences*, 109(41): 16474–16479.
- Mullainathan, Sendhil, Joshua Schwartzstein, and Andrei Shleifer. 2008. "Coarse thinking and persuasion." *The Quarterly journal of economics*, 123(2): 577–619.
- **Murciano-Goroff, Raviv.** 2017. "Missing Women in Tech: The Labor Market for Highly Skilled Software Engineers." *Working Paper*.

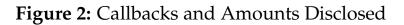
Niederle, Muriel. 2015. "Gender. Handbook of Experimental Economics."

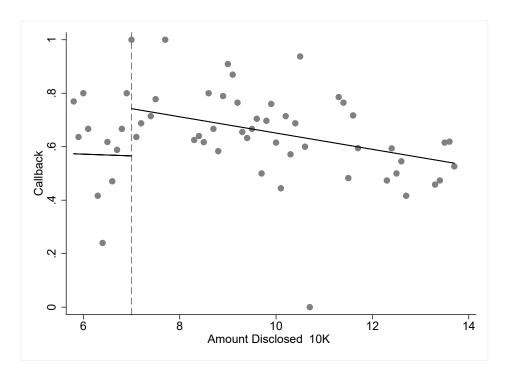
- Niederle, Muriel, and Lise Vesterlund. 2007. "Do women shy away from competition? Do men compete too much?" *The quarterly journal of economics*, 122(3): 1067–1101.
- **Oyer, Paul.** 2008. "The making of an investment banker: Macroeconomic shocks, career choice, and lifetime income." *The Journal of Finance*.
- **Oyer, Paul, Scott Schaefer, et al.** 2011. "Personnel Economics: Hiring and Incentives." *Handbook of Labor Economics*, 4: 1769–1823.
- **PayScale.** 2017. "The Salary History Question: Alternatives for Recruiters and Hiring Managers." *White Paper*.
- **Ponthieux, Sophie, and Dominique Meurs.** 2015. "Gender inequality." In *Handbook of income distribution*. Vol. 2, 981–1146. Elsevier.
- **Postel-Vinay, Fabien, and Jean-Marc Robin.** 2002. "Equilibrium wage dispersion with worker and employer heterogeneity." *Econometrica*, 70(6): 2295–2350.
- **Reuben, Ernesto, Paola Sapienza, and Luigi Zingales.** 2014. "How stereotypes impair women's careers in science." *Proceedings of the National Academy of Sciences*, 111(12): 4403–4408.
- Roussille, Nina. 2020. "The central role of the ask gap in gender pay inequality." Working Paper.
- **Rozada, Martín González, and Eduardo Levy Yeyati.** 2018. "Do women ask for lower salaries? The supply side of the gender pay gap." *Working Paper*.
- Sherman, Eliot, Raina Brands, and Gillian Ku. 2019. "The Effect of Salary History Bans on Gender-Based Disparities in Compensation: Evidence From a Field Experiment." *Working Paper*.
- Sinha, Sourav. 2019. "Salary History Ban: Gender Pay Gap and Spillover Effects." Available at SSRN 3458194.
- Solnick, Sara J. 2001. "Gender differences in the ultimatum game." *Economic Inquiry*, 39(2): 189–200.
- **Spence, Michael.** 1973. "Job Market Signaling." *The Quarterly Journal of Economics*, 87(3): 355–374.
- Sran, Gurpal, Felix Vetter, and Matthew Walsh. 2020. "Employer Responses to Pay History Inquiry Bans: Evidence from Online Job Postings, Hiring, and Pay." *Working Paper*.
- Stern, Scott. 2004. "Do scientists pay to be scientists?" Management science, 50(6): 835–853.
- Takahashi, Ana Maria, and Shingo Takahashi. 2011. "Gender salary differences in economics departments in Japan." *Economics of Education Review*, 30(6): 1306–1319.
- **Thilmany, Dawn.** 2000. "Gender based differences of performance and pay among agricultural economics faculty." *Review of Agricultural Economics*, 22(1): 23–33.
- Viscusi, W Kip. 1978. "A note on" lemons" markets with quality certification." *The Bell Journal* of Economics, 277–279.
- Westin, Alan F. 1988. "Past and future in employment testing: A socio-political overview." U. Chi. Legal F., 93.
- Wolinsky, Asher. 1983. "Prices as signals of product quality." *The review of economic studies*, 50(4): 647–658.
- Wozniak, Abigail. 2015. "Discrimination and the effects of drug testing on black employment." *Review of Economics and Statistics*, 97(3): 548–566.

Figure 1: Ban Compliance Types

		Discloses	Silent
Question Asked	Discloses	Always Discloser	Ban Complier
Salary Ques	Silent	Ban Defier	Never Discloser

Salary Question Banned





Notes: In this figure the circles present the proportion of job candidates who were recommended for a callback by the amount they disclosed. There is a line of best fit for observations in the lower 15% of the data, and one for the upper 85% of the data.

		All			Male Candida	ate	F	Female Candidate		
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	
	All	Salary Disclosed	Not Disclosed	All	Salary Disclosed	Not Disclosed	All	Disclosed	Not Disclosed	
WTP	107,101	109,922	102,978	110,814	114,873	104,882	103,388	104,970	101,075	
Outside Offer 5th %ile	88,836	93,772	81,622	92,664	99,727	82,342	85,007	87,816	80,902	
Outside Offer 50th %ile	100,044	103,635	94,796	104,235	109,783	96,127	95,852	97,486	93,465	
Outside Offer 95th %ile	112,161	113,546	110,136	117,517	119,697	114,332	106,804	107,395	105,940	
Outside Offer Range	23,325	19,774	28,514	24,853	19,970	31,990	21,797	19,579	25,038	
Offer	100,957	103,993	96,521	104,588	109,107	97,983	97,327	98,879	95,058	
Callback	0.633	0.628	0.641	0.632	0.613	0.659	0.635	0.643	0.623	
Surplus	6,144	5,929	6,458	6,226	5,766	6,899	6,061	6,091	6,016	
p(accept)	0.546	0.536	0.560	0.528	0.506	0.561	0.563	0.565	0.560	
p(accept) x Surplus	3,268	3,200	3,367	3,133	2,881	3,500	3,403	3,519	3,234	
2 Other Offers	0.530	0.526	0.536	0.560	0.549	0.575	0.501	0.503	0.498	
Offer CB	94,789	99 <i>,</i> 809	87,200	99,027	104,658	89,707	90,517	94,558	84,932	
Observations	2048	1216	832	1024	608	416	1024	608	416	

Table 1: Candidate Summary Statistics

Notes: Each of our 256 recruiters evaluated eight candidates for a total of 2048 candidate level observations. Willingness-to-Pay (WTP) is the maximum a recruiter stated we should pay a particular candidate. Outside offer *X* %ile is their answer to the question "what offer would this candidate accept with *X* probability". Outside offer range is Outside Offer 95th %ile - Outside Offer 5th %ile. Offer is the take-it-or-leave-it-offer the recruiter suggested. Callback is whether the recruiter suggested the candidate be interviewed. Surplus is WTP minus Offer. p(accept) is the probability that a candidate would accept the specific take-it-or-leave it offer as approximated by fitting a logistic function through the *X* %ile answers. p(accept) *x* Surplus is the expected surplus from a certain take-it-of-leave it offer. 2 Other Offers is whether the recruiter believes this candidate would have 2 or more other offers (as opposed to 1 or 0). Offer | CB is the amount of the offer conditional on the recruiter recommending the candidate for a callback.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	WTP	WTP	Outside Offer 50th %ile	Outside Offer 50th %ile	Offer	Offer	Outside Offer Range	Outside Offer Range
Salary Disclosed	0.68***	0.97***	0.84***	1.30***	0.73***	1.08***	-0.92***	-1.25**
Female x Disclosed	(0.14)	(0.16) -0.57***	(0.14)	(0.15) -0.92***	(0.14)	(0.16) -0.69***	(0.24)	(0.40) 0.67+
		(0.13)		(0.11)		(0.12)		(0.37)
Female Disclosure Effect:								
Total		0.40		0.38		0.39		-0.59
<i>p-value</i> Mean Non-Dislosers:		0.01		0.01		0.01		0.00
All	10.30	10.30	9.48	9.48	9.65	9.65	2.85	2.85
Male	10.49	10.49	9.61	9.61	9.80	9.80	3.20	3.20
Female	10.11	10.11	9.35	9.35	9.51	9.51	2.50	2.50
R ²	0.18	0.18	0.25	0.26	0.20	0.21	0.01	0.02
Observations	2048	2048	2048	2048	2048	2048	2048	2048

Table 2: Average Effect of Disclosing Salary

Notes: All models include recruiter controls and candidate fixed effects. This table shows estimates from versions of Equation 5. Dependent variables are listed in the column header and explained in notes to Table 1. Outcomes measured in dollars (e.g. WTP, Offer) are in \$10K increments. Robust standard errors are clustered at the recruiter level. + p < 0.10 * p < 0.05 * * p < 0.010 * * p < 0.001

	(1)	(2)	(3)	(4)
	WTP	Outside Offer	Offer	Outside Offer
	VV 11°	50th %ile	Oller	Range
Disclosed 25th %ile Salary	-0.08	-0.04	-0.04	-0.52
	(0.14)	(0.13)	(0.12)	(0.34)
Disclosed 75th %ile Salary	1.11***	1.24***	1.18***	-0.37
	(0.15)	(0.14)	(0.13)	(0.29)
Mean Non-Disclosers	10.30	9.48	9.65	2.85
R ²	0.34	0.46	0.38	0.01
Observations	2048	2048	2048	2048

Table 3: Average Effect of Disclosing a High versus Low Salary

Notes: All models include recruiter and spillover controls and both candidate and sub-treatment fixed effects as described in the text. Dependent variables are listed in the column header and explained in notes to Table 1. Outcomes measured in dollars (e.g. WTP, Offer) are in \$10K increments. Disclosed Xth %tile Salary means a candidate disclosed a salary at the Xth percentile within their specific firm. The omitted category is candidates who did not disclose a salary. Robust standard errors are clustered at the recruiter level. + p < 0.10 * p < 0.05 * p < 0.010 * p < 0.001

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	WTP	WTP	Outside Offer 50th %ile	Outside Offer 50th %ile	Offer	Offer	Outside Offer Range	Outside Offe Range
Disclosed x Amt Disclosed	0.65***	0.67***	0.77***	0.78***	0.68***	0.69***	-0.03	0.03
	(0.06)	(0.06)	(0.05)	(0.05)	(0.05)	(0.06)	(0.12)	(0.18)
Salary Disclosed	-5.76***	-6.17***	-6.82***	-7.11***	-6.04***	-6.33***	-0.19	-1.12
-	(0.56)	(0.61)	(0.49)	(0.54)	(0.53)	(0.56)	(0.92)	(1.68)
Female x Disclosed x Amt Disclosed		0.01		-0.00		0.02		-0.03
		(0.05)		(0.04)		(0.04)		(0.14)
Female x Disclosed		0.34		0.27		0.19		1.02
		(0.47)		(0.39)		(0.41)		(1.49)
Female Amount Disclosed Slope:								
Total		0.68		0.78		0.71		-0.00
p-value		0.00		0.00		0.00		0.99
Mean Non-Dislosers:								
All	10.30	10.30	9.48	9.48	9.65	9.65	2.85	2.85
Male	10.49	10.49	9.61	9.61	9.80	9.80	3.20	3.20
Female	10.11	10.11	9.35	9.35	9.51	9.51	2.50	2.50
R ²	0.38	0.38	0.54	0.54	0.44	0.44	0.01	0.01
K								

Table 4: Average Effect of Disclosing by Salary Amount

Notes: All models include recruiter and spillover controls and both candidate and sub-treatment fixed effects as described in the text. This table shows estimates from versions of Equation 5 that include interactions with gender. Dependent variables are listed in the column header and explained in notes to Table 1. Salary amounts and outcomes measured in dollars (e.g. WTP, Offer) are in \$10K increments. Robust standard errors are clustered at the recruiter level. + p < 0.10 * p < 0.05 ** p < 0.010 *** p < 0.001

	(1)	(2)	(3)	(4)	(5)
	WTP	Outside Offer	Offer	O.O.	p WTP =
	VV 11	50th %ile	Ollei	Range	Outside Offer
+\$10k from Firm	0.64***	0.77***	0.68***	-0.01	0.00
	(0.07)	(0.06)	(0.06)	(0.20)	
+\$10k from Male	0.42***	0.62***	0.48***	-0.39	0.00
	(0.09)	(0.07)	(0.09)	(0.26)	
+\$10k within Firm	0.70***	0.79***	0.73***	0.04	0.06
	(0.06)	(0.05)	(0.05)	(0.06)	
<i>p</i> F-M	0.00	0.01	0.00	0.12	
p F-W	0.27	0.69	0.28	0.77	
<i>p</i> M-W	0.00	0.01	0.00	0.09	
2					
\mathbb{R}^2	0.38	0.54	0.44	0.01	
Observations	2048	2048	2048	2048	

Table 5: Effect of An Extra Dollar Decomposed

Notes: All models include recruiter and spillover controls and both candidate and sub-treatment fixed effects as described in the text. This table shows estimates from Equation 6 which decomposes additional dollars of salary disclosure into a firm-specific offset for the candidate's employer ("+\$10k from Firm," some firms pay higher or lower to everyone on average); a gender offset ("+\$10k from Male", which mimics real-world gender gaps); and from having a higher or lower salary within the current firm's distribution ("+\$10k within Firm", note this also is combined with some random noise that was included in the salaries.). Dependent variables are listed in the column header and explained in notes to Table 1. Outcomes measured in dollars (e.g. WTP, Offer) are in \$10K increments. Column 5 reports the *p*-value on a test of whether the coefficient on WTP in Column (1) = the coefficient on Outside Offer 50th %ile in Column (2). *p*-values for comparisons of coefficients within the same model are provided in the 2nd panel, where for example *p* F-M is the p-value testing that the coefficient from "+\$1 from Firm" = the coefficient on "+\$1 from Male". Robust standard errors are clustered at the recruiter level. + *p* < 0.10 * *p* < 0.05 ** *p* < 0.010 *** *p* < 0.001

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Callback	Callback	Surplus	Surplus	p(accept)	p(accept)	p(accept) x Surplus	p(accept) x Surplus
Salary Disclosed	-0.01	-0.04	-0.05	-0.11*	-0.01	-0.04	-0.01	-0.06+
	(0.02)	(0.03)	(0.04)	(0.04)	(0.02)	(0.03)	(0.03)	(0.03)
Female x Disclosed		0.07 +		0.12**		0.06**		0.09**
		(0.04)		(0.04)		(0.02)		(0.03)
Female Disclosure Effect: Total		0.02		0.01		0.01		0.03
<i>p-value</i> Mean Non-Dislosers:		0.48		0.80		0.57		0.30
All	0.64	0.64	0.65	0.65	0.56	0.56	0.34	0.34
Male	0.66	0.66	0.69	0.69	0.56	0.56	0.35	0.35
Female	0.62	0.62	0.60	0.60	0.56	0.56	0.32	0.32
R ²	0.03	0.03	0.01	0.01	0.03	0.04	0.02	0.03
Observations	2048	2048	2048	2048	2048	2048	2048	2048

Table 6: Average Effect of Disclosing for Callback Outcomes

Notes: All models include recruiter controls and candidate fixed effects. This table shows estimates from versions of Equation 5. Dependent variables are listed in the column header and explained in notes to Table 1. Outcomes measured in dollars (e.g. surplus) are in \$10K increments. Robust standard errors are clustered at the recruiter level. + p < 0.10 * p < 0.05 * p < 0.010 * p < 0.001

Table 7: Average Effect of Disclosing for Callback OutcomesSalary > \$70,000

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Callback	Callback	Surplus	Surplus	p(accept)	p(accept)	p(accept) x Surplus	p(accept) x Surplus
Salary Disclosed	-0.00	-0.03	-0.08*	-0.12*	-0.02	-0.04	-0.04	-0.06*
	(0.03)	(0.03)	(0.04)	(0.05)	(0.02)	(0.03)	(0.03)	(0.03)
Female x Disclosed		0.07		0.09*		0.05 +		0.06+
		(0.04)		(0.04)		(0.02)		(0.03)
Female Disclosure Effect:								
Total		0.03		-0.03		0.00		-0.01
p-value		0.32		0.42		0.92		0.83
, Mean Non-Dislosers:								
All	0.64	0.64	0.65	0.65	0.56	0.56	0.34	0.34
Male	0.66	0.66	0.69	0.69	0.56	0.56	0.35	0.35
Female	0.62	0.62	0.60	0.60	0.56	0.56	0.32	0.32
R ²	0.02	0.02	0.01	0.01	0.03	0.03	0.01	0.01
Observations	1849	1849	1849	1849	1849	1849	1849	1849

Notes: All models include recruiter controls and candidate fixed effects. This table shows estimates from versions of Equation 5 and mimics Table 6; the sample is restricted to candidates who did not disclose and those who disclosed more than \$70,000. See Section 7.4.1 and Figure 2 for explanation of the threshold. Dependent variables are listed in the column header and explained in notes to Table 1. Outcomes measured in dollars (e.g. Surplus) are in \$10K increments. Robust standard errors are clustered at the recruiter level. + p < 0.10 * p < 0.05 * * p < 0.010 * * * p < 0.001

	(1)	(2)	(3)	(4)
	Callback	Surplus	p(accept)	p(accept) x
	Caliback	Surpius	placept)	Surplus
+\$10k from Firm	-0.02	-0.04+	-0.02	-0.04**
	(0.02)	(0.02)	(0.01)	(0.01)
+\$10k from Male	-0.05+	-0.07*	-0.03	-0.05*
	(0.03)	(0.03)	(0.02)	(0.02)
+\$10k within Firm	0.00	-0.03	-0.01	-0.03+
	(0.02)	(0.02)	(0.01)	(0.02)
<i>p</i> F-M	0.27	0.25	0.76	0.76
p F-W	0.28	0.68	0.46	0.36
<i>p</i> M-W	0.07	0.19	0.37	0.36
R ²	0.04	0.04	0.07	0.06
Observations	2048	2048	2048	2048

Table 8: Effect of An Extra Dollar Decomposed for Callback Outcomes

Notes:All models include recruiter and spillover controls and both candidate and sub-treatment fixed effects as described in the text. This table shows estimates from Equation 6 which decomposes additional dollars of salary disclosure into a firm-specific offset for the candidate's employer ("+\$10k from Firm," some firms pay higher or lower to everyone on average); a gender offset ("+\$10k from Male", which mimics real-world gender gaps); and from having a higher or lower salary within the current firm's distribution ("+\$10k within Firm", note this also is combined with some random noise that was included in the salaries.). Dependent variables are listed in the column header and explained in notes to Table 1. Outcomes measured in dollars (e.g. Surplus) are in \$10K increments. *p*-values for comparisons of coefficients within the same model are provided in the 2nd panel, where for example *p* F-M is the p-value testing that the coefficient from "+\$1 from Firm" = the coefficient on "+\$1 from Male". Robust standard errors are clustered at the recruiter level. + *p* < 0.10 * *p* < 0.05 ** *p* < 0.010 *** *p* < 0.001

Table 9: Average Effect of Disclosing by Salary Amount for CallbackOutcomes

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Callback	Callback	Surplus	Surplus	p(accept)	p(accept)	p(accept) x Surplus	p(accept) x Surplus
Salary Disclosed	0.13	0.18	0.28	0.17	0.15	0.12	0.33*	0.26+
	(0.14)	(0.18)	(0.17)	(0.21)	(0.12)	(0.14)	(0.13)	(0.15)
Female x Disclosed		-0.17		0.15		0.03		0.10
		(0.17)		(0.19)		(0.10)		(0.12)
Disclosed x Amt Disclosed	-0.01	-0.02	-0.03*	-0.03	-0.02	-0.02	-0.03**	-0.03*
	(0.01)	(0.02)	(0.02)	(0.02)	(0.01)	(0.01)	(0.01)	(0.01)
Female x Disclosed x Amt Disclosed		0.02		-0.01		-0.00		-0.01
		(0.02)		(0.02)		(0.01)		(0.01)
Female Amount Disclosed Slope:								
Total		0.01		-0.04		-0.02		-0.04
p-value		0.71		0.07		0.23		0.01
Mean Non-Dislosers:		0.71		0.07		0.20		0.01
All	0.64	0.64	0.65	0.65	0.56	0.56	0.34	0.34
Male	0.66	0.66	0.69	0.69	0.56	0.56	0.35	0.35
Female	0.62	0.62	0.60	0.60	0.56	0.56	0.32	0.32
R ²	0.04	0.04	0.04	0.04	0.07	0.07	0.06	0.06
Observations	2048	2048	2048	2048	2048	2048	2048	2048

Notes: All models include recruiter and spillover controls and both candidate and sub-treatment fixed effects as described in the text. This table shows estimates from versions of Equation 5. Dependent variables are listed in the column header. Salary amounts and outcomes measured in dollars (e.g. Surplus) are in \$10K increments. Robust standard errors are clustered at the recruiter level. + p < 0.10 * p < 0.05 ** p < 0.010 *** p < 0.001

Table 10: Average Effect of Disclosing by Salary Amount for CallbackOutcomes Salary > \$70,000

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Callback	Callback	Surplus	Surplus	p(accept)	p(accept)	p(accept) x Surplus	p(accept) x Surplus
Salary Disclosed	0.43*	0.56*	0.22	0.29	0.30+	0.28	0.37*	0.42*
	(0.18)	(0.22)	(0.21)	(0.26)	(0.16)	(0.17)	(0.15)	(0.16)
Female x Disclosed		-0.34		-0.19		0.03		-0.09
		(0.21)		(0.24)		(0.13)		(0.14)
Disclosed x Amt Disclosed	-0.04*	-0.05*	-0.03	-0.04	-0.03+	-0.03+	-0.04**	-0.04**
	(0.02)	(0.02)	(0.02)	(0.02)	(0.02)	(0.02)	(0.01)	(0.02)
Female x Disclosed x Amt Disclosed		0.04+		0.02		-0.00		0.01
		(0.02)		(0.02)		(0.01)		(0.01)
Female Amount Disclosed Slope:								
Total		-0.01		-0.02		-0.03		-0.03
p-value		0.46		0.44		0.08		0.04
, Mean Non-Dislosers:								
All	0.64	0.64	0.65	0.65	0.56	0.56	0.34	0.34
Male	0.66	0.66	0.69	0.69	0.56	0.56	0.35	0.35
Female	0.62	0.62	0.60	0.60	0.56	0.56	0.32	0.32
R ²	0.04	0.04	0.04	0.04	0.07	0.07	0.04	0.04
Observations	1849	1849	1849	1849	1849	1849	1849	1849

Notes: The sample in this table is restricted to those that did not disclose a salary and those that disclosed a salary above \$70,000. All models include recruiter controls and candidate fixed effects. This table shows estimates from versions of Equation 5. Dependent variables are listed in the column header and explained in notes to Table 1. Salary amounts and outcomes measured in dollars (e.g. surplus) are in \$10K increments. Robust standard errors are clustered at the recruiter level. + p < 0.10 * p < 0.05 ** p < 0.010 *** p < 0.001

	Type of Cel	lls	Proporti	ion of Cell tha	t is an		Mean Disclosure Δ		
Gender	High/Low Wage Firm	High/Low Salary vs Peers @Firm	Always Discloser	Never Discloser	Ban Complier	% of Sample	Callback Prob	TIOLI Offer	
Female	High	High	0.28	0.24	0.48	11.6	+0.02	+\$21.7K***	
Female	High	Low	0.23	0.21	0.55	11.1	+0.06	+\$2.7K	
Female	Low	High	0.18	0.19	0.63	15.4	+0.03	+\$834	
Female	Low	Low	0.20	0.15	0.65	11.5	-0.04	-\$13.3K***	
Male	High	High	0.33	0.23	0.43	15.0	-0.15**	+\$28.3K***	
Male	High	Low	0.42	0.15	0.43	11.4	-0.003	+\$12.7K***	
Male	Low	High	0.31	0.19	0.49	11.6	+0.05	+\$7.38K***	
Male	Low	Low	0.31	0.15	0.54	12.2	-0.05	-\$9.79K***	

Table 11: Complier Type Cells

Notes: This table aggregates data from our survey of 1,006 US job seekers (middle columns) and our experiment (final two columns). It reports the distribution of compliance types by gender, whether a person works at a high wage firm, and whether they self-report having a salary above the median within their firm. The final columns report the average changes in outcomes from candidate disclosures in our experiment. An "Always Discloser" always reports their salary whether prompted or not, a "Never Discloser" never reports, and a "Ban Complier" will disclose if prompted.

	Women	Men	Ratio
No Ban	.64	.63	1
	(.023)	(.023)	(.045)
Ban	.64	.63	1
	(.029)	(.024)	(.05)
Ban-No Ban	0034	0027	00096
	(.037)	(.033)	(.067)
<i>p</i> -value	.93	.93	.99

Table 12: Effect of Bans (Full or Partial) on Callbacks

Callbacks

Notes: This table shows the effect of a salary history ban on whether a candidate was recommended for a callback. Under both a full or partial ban the callback decision is always made when only the information from the application is available either with a prompt under No Ban, or without a prompt under Ban. Standard errors are robust.

Table 13: Full vs. Partial Ban: Effect on Annual Salaries, Conditional on Callbacks

Salam | Callbacks (Dartial Ban)

Sutury Cattoach	(Fuii Duni)			Satury Cuttback			
	Women	Men	Ratio		Women	Men	Ra
No Ban	101824.98	112220.58	0.91	No Ban	101824.98	112220.58	0.9
	(1111.61)	(1315.39)	(0.01)		(1111.61)	(1315.39)	(0.0
Ban	100377.55	103921.13	0.97	Ban	97966.75	106061.10	0.9
	(1576.33)	(1349.36)	(0.01)		(319.37)	(478.83)	(0.0
Ban-No Ban	-1447.43	-8299.45	0.06	Ban-No Ban	-3858.23	-6159.48	0.0
	(1928.86)	(1884.42)	(0.01)		(1156.58)	(1399.83)	(0.0
<i>p</i> -value	0.45	0.00	0.00	<i>p</i> -value	0.00	0.00	0.1

Salary | Callbacks (Full Ban)

Notes: This table shows the effect of a salary history ban on the salary offer when a candidate was recommended for a callback. The left panel shows this for a full ban and the right panel is for a partial ban. A "Full Ban" means a ban where salary history may not be asked at any stage in the hiring process. A "Partial Ban" means a ban of prompting job candidates to disclose their salary history on the job application, but being able to seek salary information at a later stage in the hiring process. Standard errors are robust.

Appendix: For Online Publication Only

A Labor Market: Software Engineering

We tasked our recruiting workforce with screening applicants for a software engineer position. The software sector is an ideal labor market for studying the effects of salary disclosures and bans on asking for salary history. The market for software jobs features several particularly attractive features for this study.¹

First, technical roles exhibit persistent gender wage and employment differences that span multiple decades (Blau, Brummund and Liu, 2013; Goldin et al., 2017). Only 19% of computer science degrees are held by women, and one-third of workers in the technology sector in Silicon Valley are female.² Given the high wage and employment growth in this sector, technology may be a growing source of income inequality overall (Krueger, 1993; Acemoglu and Autor, 2011). Second, the technology sector features well-documented labor shortages and high levels of competition between employers for qualified workers. Technology executives regularly lobby Congress for expansions to the H1-B visa program to address the undersupply of software developers. Firms in this sector are generally interested in hiring multiple qualified candidates whenever possible. As the H1-B lobbying shows, hiring is limited not by demand, but by the supply of qualified workers. As a result, we can measure how salary disclosures and bans of prompting disclosures affect the likelihood of a candidate being called back, in addition to how the composition of selected candidates and salaries change.

Second, by choosing this industry, we bias our study toward finding smaller differences between experimental variations. Labor shortages should erode the effect of gender and past salary on evaluations of our job candidates. With strong competition for qualified candidates, there is likely to be less taste-based discrimination (Becker, 1957). This might lead salary-history bans to be less effective in this industry. Behavioral-economics phenomena such as "framing" and "anchoring" are often used to motivate why salary disclosures can be harmful and why salary-history bans might reduce wage gaps. Effects in other, less-competitive sectors may be stronger.

B Details of Randomization Procedure

Our randomization procedure was sequential, proceeded in batches, and was designed to address covariate balance through re-randomization. For recruiters who were invited, accepted, and met our prescreening qualifications (i.e., signed a nondisclosure agreement and possessed relevant experience), the recruiters' demographics were manually coded.³ We merged the coded demographics data with data about the recruiter's prior work experiences and posted wage rate.

¹We chose to examine the market for engineers with moderate experience so that our candidates had a previous wage history that could (or could not) be disclosed.

²See https://nces.ed.gov/programs/digest/d18/tables/dt18_325.35.asp and https: //www.bloomberg.com/news/articles/2019-02-13/silicon-valley-is-using-trade-s ecrets-to-hide-its-race-problem

³For our full sample of recruiters, the recruiter's self-reported gender matched our manually coded gender 99% of the time. The recruiter's self-report of identifying as black matched our manual coding of this variable 92% of the time, while a recruiter's self-report of identifying as white matched our manual coding 87% of the time.

Before sending out the experimental materials for recruiters' feedback, we performed a covariate balance check (described below). If our covariate balance test passed, we would send the experimental materials to the recruiters. If the balance checks failed, we re-randomized the current batch (previous batches had already been sent to recruiters, who had begun work on them, so they could not be re-randomized).

Our balance test checked for equality of the average of the following covariates across treatment arms. The covariates were as follows: 1) race (dummy variables for white and black), 2) gender, 3) the recruiter's advertised hourly rate, and 4) a dummy variable for whether the recruiter had previously logged hours on the website we used to hire them.

We tested for equality of these means across all treatment groups (a single test per variable for equality across all treatment arms). In addition, we tested for pairwise equality across all treatment arms. For assignments where these tests' *p*-values were less than 0.2, we re-randomized. We also randomized if the pairwise comparison for any two subtreatments was less than 0.05.

The sequential balance checks were cumulative. The tests above included observations for all prior assignments, including the current batch. However, the current batch was the only batch that could be potentially adjusted if re-randomization proved to be necessary. Batches were processed approximately once per week, so that recruiters would not have to wait long after accepting our offer to begin work.

B.1 Recruiter Characteristics Balance

Our study randomized the salary-history prompt, proportion disclosing, and distribution of amounts disclosed at the recruiter level. Prior research suggests that hiring decisions differ according to managers' characteristics.⁴ As such, we implemented a randomization procedure to guarantee covariate balance on recruiter characteristics such as race and gender across recruiter-level variations. This effectively implemented stratified randomization, guaranteeing that (for example) male recruiters were not overassigned to one particular experimental arm by accident.

Table G2 shows that our stratification procedure succeeded; the recruiter demographics are balanced across whether the recruiter was shown applications with a prompt or not, and whether that recruiter was shown zero, four, or eight candidates who disclosed. Almost none of the mean differences between our main experimental variations approach traditional levels of statistical significance.⁵

⁴For example, Giuliano, Levine and Leonard (2009) report that nonblack managers hire more white workers and fewer black workers. (Dee, 2005) finds that educators evaluate students of the opposite gender more harshly.

⁵The proportion of screeners who are black is 28% for those shown four disclosures, while it is 20% for those shown zero disclosures, a comparison which has a one-sided *t*-test of Pr(T > t) = 0.0956. The proportion who had been asked for salary input before is 100% for those shown zero disclosures, while it is 97% for those shown four salary disclosures. The comparison has a one-sided *t*-test of Pr(T > t) = 0.0919, a difference that is statistically significant but likely not economically significant. We randomized three things at the recruiter level: 1) prompt, 2) proportion disclosed, and 3) distribution of amounts disclosed. The interaction of those three variations results in 22 distinct recruiter-level subtreatments. In Table G2, we show the mean of the recruiter characteristics across these subtreatments. There are a total of 546 two-way comparisons, and of these, 16% are statistically significant at traditional levels. As such, we include controls for screen characteristics in our models.

C Details of Recruiter Selection

The platform we used allowed workers to either be paid an hourly rate or to negotiate a pay-by-task contract. Each recruiter's profile includes an hourly rate suggested by the recruiter. We offered to pay our subjects the hourly rate posted on their profile. We also offered a bonus contract designed to align their interests with the firm's as they made decisions. The incentive scheme assessed the recruiters' decision based on candidates' outcomes, and it rewarded better decision-making according to a formula.⁶ Such bonus systems are common in real-world HR work.

All recruiters worked remotely and corresponded with us directly over the Internet. Each qualified recruiter was sent the materials containing a set of applications to review and an online application for qualitative and quantitative assessments of each candidate, along with commentary. Recruiters were also sent a description of the firm and the hiring needs for the opening.

Each recruiter was required to sign a nondisclosure agreement, a common practice in real-world recruitment outsourcing in order to protect firm and candidate confidentiality. All these materials are available in the Experimental Materials Appendix. We did not directly tell any recruiters that they were part of a larger recruiting workforce containing peers, but our instructions did reference the firm's other HR staff. The NDA also helped to address the possibility that recruiters would discover each other through circumstance and discuss the assignment. All recruiters signed the NDA, although some felt it was unnecessary because it was covered by the platform's terms of service.

To be eligible for an invitation into our workforce, recruiters on the platform had to be listed as independent (rather than affiliated with an agency),⁷ based in the United States according to their profile,⁸ and had to have worked previously in real-world recruiting roles for office jobs.

We searched using key words such as "recruiter," "sourcing," "talent acquisition," "staffing," and "human resources." We did this in two waves. Wave 1 took place in the summer of 2018, while wave 2 was executed in late 2019. Over both waves, a list of approximately 20,000 possible recruiters was identified on key words, then we examined a random sample of approximately 5,000, and research assistants marked about 1,750 recruiters as qualified, by checking the recruiter's profile for prior real-world experience in hiring or recruiting for nonmanual work. We then invited each qualified recruiter charging less than or equal to \$100 an hour.⁹ Approximately 400 wrote

⁷We did not hire agencies, in order to avoid the possibility of recruiters in different treatment arms having discussions among each other.

⁸We focused on US-based recruiters who would be familiar with the qualifications of US-based candidates.

⁹The recruiters all indicated interest in HR or hiring through the key words they put in their profile. We also asked each invited recruiter for a resume or LinkedIn profile. Before officially having them start the project, we checked these resumes/profiles for hiring experience. If the experience wasn't clear, we offered them the chance to clarify by

⁶The bonus formula for a hired candidate was

⁵x(technicalscore + innovationscore + leadershipscore) - Salary/100000.

Each part of the technical/innovation/leadership score took integer values between one and three, and recruiters were told these scores would be assigned after a worker had been on the job for four weeks (details available in Section H). Recruiters also lost \$5 for each candidate they suggested calling back who was not interested in being hired. This could result in a negative bonus, so recruiters were told that at worst they would earn a \$0 bonus. These bonuses were eventually paid. However, since there were no actual candidates or firms, we simulated the outcomes of callbacks, salary offers/acceptances, and performance from data from similar real-world firms. Recruiters were led to believe that the bonus was paid based on real-world outcomes, including an assessment of the candidate four weeks after he or she started the job. All bonuses were paid between 30 and 45 days after we received feedback from a recruiter, which we believed was a reasonable time period for a candidate to be hired and given an initial assessment of value by a firm.

back in response to our inquiry to accept the job offer within the time frame of our experiment. Most of the remainder did not write back at all, or wrote back after the experiment was completed.

These job requirements are typical for recruiting. The BLS's occupational data suggest that human resource work is mid-skill work, requiring a bachelor's degree but no related work experience or prior on-the-job training.¹⁰ According to the BLS, our requirement of prior experience for recruiters is actually more stringent than a typical requisition for a recruiter. Over 70% of our subjects reported having more than three years' prior experience, and 98% stated that they had provided salary input in prior recruiting assignments. We did not require prior experience specifically in recruiting software engineers. However, prior experience in software engineering recruiting is not necessary for a recruiting job at many tech companies, as hiring for high-skill jobs is quite similar across many sectors (Adler, 2020).

D Details of Creating Candidates

For first names, we used the top four given names of each sex for Americans according to the Social Security Administration (making job candidates between 24–27 years old at the time we began our experiment).¹¹ We blacked out the last name so that recruiters could not try to contact our candidates or look them up online (Acquisti and Fong, 2015); we also encouraged recruiters to make decisions based on the application materials rather than investigating them online.

Each candidate was assigned a bachelor's degree in computer science from universities ranked from third to ninth in the country in computer engineering by *US News and World Report*.¹² We excluded the top two universities (MIT and Berkeley) to avoid the possibility that the top institutions might have some special cache, since variation in school quality was not one of primary variations of interest for the experiment.

Previous firms were chosen from the top firms that hire software engineers.¹³ To ascertain previous salaries, we matched these firms with salaries reported on PayScale.com.¹⁴ PayScale

¹²There are in fact nine schools ranked between third and ninth as a result of ties. They are: Carnegie Mellon, the University of Illinois Urbana-Champaign (UIUC), Georgia Tech, the University of Michigan, University of Texas at Austin, Cornell University, Cal Tech, the University of Washington, and Purdue University. We randomly selected from the three schools tied for ninth place, so that our final applicants did not attend Purdue. See: https://www.usnews.com/best-graduate-schools/top-engineering-schools/computer-engineering-rankings.

asking them to tell us about their hiring experience. If this answer implied that a firm would be interested in hiring this person for this role based on their response, then we proceeded. Approximately 40 individuals who responded to our initial inquiry were ultimately not sent experimental materials, mostly because they had insufficient experience with hiring/recruiting/screening.

¹⁰https://www.bls.gov/ooh/business-and-financial/human-resources-specialists
.htm#tab-1

¹¹Male names were Andrew, Tyler, Joshua and Christopher. Female names were Jessica, Emily, Samantha, and Sarah. See https://www.ssa.gov/oact/babynames/top5names.html. We excluded the name "Ashley" as it could be interpreted as being either male or female.

¹³See https://www.techrepublic.com/article/the-10-companies-hiring-more-soft ware-engineers-than-anyone-else-in-silicon-valley/ and https://www.monster.com/ career-advice/article/top-tech-employers-job-listings.

¹⁴We also verified that PayScale.com's estimates were comparable to those on Glassdoor.com, a similar website collecting salary data: for example, https://www.payscale.com/research/US/Job=Software_E ngineer/Salary/3f79787f/Amazon.com-Inc-Seattle-WA and https://www.glassdoor.

provides very granular data indexed by company, job roles, city and level of experience. We obtained the 25th, 50th and 75th percentile of salaries for software engineers with one-to-three years of experience in each firm's headquarter city.¹⁵

Each candidate biography required a realistic salary that could be disclosed when assigned to disclosure treatments. To approximate realistic gender gaps in salaries, we analyzed data from the 2015 American Community Survey (ACS).¹⁶

Our goal is to adjust the firm-city specific salaries from PayScale to create plausible male and female salaries for all candidate biographies. We adjust the PayScale salaries for men at each firm by multiplying the appropriate salary by 1.05. Then we multiply the result by 0.80 to get the estimated female salaries at the same firm, location, and job. We derived these estimates from our analysis of the ACS data.¹⁷

The salaries reported on our job applications use these numbers, with a few additional adjustments: we added a small amount of noise¹⁸ and rounded to the nearest \$1,000. The noise and rounding produced only trivial changes to the distribution of salaries. However, it guaranteed that the "roundness" of disclosed salary numbers was randomly assigned and uncorrelated with a candidate's gender, current employer, or other characteristics. Prior research suggests that round numbers are received differently in negotiation (Mason et al., 2013).

Each applicant had one job after graduation before his or her current job, as well as a college internship. Two jobs since graduation is typical, considering our candidates were in the full-time workforce for four to five years by the time of their applications.¹⁹ We injected small amounts of random variation in the start date and duration of the first job. This was in order to create realistic variation across candidates so they did not all contain identical dates. The postcollege job started shortly after college graduation and had a total tenure of between 6 and 17 months (randomly selected). The duration of the current job varied by when the recruiter viewed the applicant's materials, but all the current jobs started between February 2014 and November 2015. The applications also listed additional skills, achievements, and coursework modeled after the resumes of real software engineers.

¹⁵For IBM, which had no software engineer salary data in its headquarters of Armonk, N.Y., we instead used salaries from its other major campus in San Jose, Calif.

¹⁶The actual wage gap is difficult to compute and is beyond the scope of this paper. Publicly available salary data about specific firms—including the sources we used above (Glassdoor and PayScale) and all others we consulted—do not contain gender-specific wage values.

¹⁷We restrict the ACS data to individuals with a bachelor's degree (only), who are employed either in computer occupations (ACS Occupation Codes 10XX and 11XX) or specifically as computer software engineers (ACS Occupation Code 1020). Note that our PayScale data combine data for men and women. On average, in the ACS, men in both computer and specifically software engineer occupations make 1.05 times the overall average. For computer occupations, women make on average 0.81 times what men make; for software developer occupations, women make on average 0.78 times what men make.

¹⁸This draws from a uniform random distribution from -\$2,000 to +\$2,000 in \$1K increments.

¹⁹According to the BLS, median job tenure for those aged 20 to 24 is 1.3 years, and for those aged 25 to 34 it is 2.8 years (https://www.bls.gov/news.release/tenure.t01.htm).

com/Salary/Amazon-Software-Engineer-Salaries-E6036_D_K07,24.htm. The distribution of base salaries reported to these types of websites is quite similar to those reported to the U.S. census. For example, Glassdoor.com has benchmarked its salary data against census data and published the results several times, and they are remarkably similar for base pay (Glassdoor, 2019).

D.1 Candidate Characteristics Balance

We have full control of all the attributes of the job candidates, including whether they disclose, so we made sure to balance our candidates on attributes we were not primarily interested in. For example, the average year of graduation was 2013, and the proportion currently working at Amazon is 6% for candidates who don't disclose as well as for those who do disclose, as shown in Table G1.

E Spillover Controls

We also include spillover terms that take into account the disclosures of other candidates who were included in the same packet of eight sent to the recruiters. In specifications that do not include the amount of salary disclosed, e.g. Equation 5, to account for potential spillovers we control for the number of other applicants in the packets whose salaries *are* disclosed (this can be either 0, 3, 4, or 7). Each line of the data is a single job candidate. In specifications that do include the amount disclosed, e.g., Equation 6, we further control for the average of all the other salaries disclosed amongst the eight, excluding the job candidate's own, and fixed effects for the subtreatment the packet was assigned.²⁰

E.1 Prompted vs. Unprompted Disclosures

Our approach in this section allows for the possibility that prompts affect not only which candidates disclose but also how employers interpret those disclosures. Before combining our survey data with our field experiment, first we examine the asymmetric effects of *unprompted* disclosures within our field experiment. We measure this by estimating Equation 5 with prompt interactions, and we present results in Table G10. For most of our outcomes, we find relatively small, statistically insignificant differences. The only statistically significant effect is that the range of outside offers contracts more under prompted disclosure than it does under unprompted disclosure. Although the differences are insignificant for our other outcomes, the direction of the effects tells a common story: prompted disclosures matter more than unprompted ones.

This may have a simple explanation: Prompted disclosures are more noticeable than unprompted ones. The prompt may direct visual and strategic attention to the disclosure behavior. By contrast, unprompted disclosures in our experiment require a recruiter to read the "additional information" section, notice the disclosure, and realize its significance. Although this may be an artifact of our experimental setup, similar results may happen more broadly. An unprompted disclosure to the wrong person—an interviewer instead of an HR person, or an HR person instead of the boss—may not reach the key decision maker. Additionally, the existence of a prompt for salary history on the job application may signal that an employer values this information, and thus our recruiters may rely on it more.

We also investigate whether unprompted disclosure has gendered effects. Given gender stereotypes and cultural expectations, one may wonder whether unprompted disclosures by women might

²⁰For example, if the first line of the data is for Jessica and she discloses \$105,000, and three other people in the packet disclose \$90,000, \$97,000, and \$103,000, then we include the average of those three disclosures (\$97,000) as a control variable.

evoke negative reactions. We can generally rule out large gendered effects.

This subsection has discussed the effect of prompted disclosure versus unprompted disclosure on an individual job seeker's outcomes. However, to simulate the effects of a ban on gender inequality, we want to estimate the effect of a ban while taking into account the proportion of job seekers who comply with the ban, always disclose, or never disclose.

F Recruiter Knowledge of Average Market Wages

One potential alternative explanation for our results is that recruiters simply misjudged the average level of market wages for this job. Our subjects may have believed that silent workers earned market-average wages but misjudged average-market pay levels for software engineers. Our candidates' disclosure amounts were based on third-party data about true, accurate market levels, and our recruiter subjects were experienced professionals. Insofar as they were not, they could estimate market levels using the same publicly available tools. In fact, we administered a brief questionnaire to the recruiters after they completed the main task, and we found that when recruiters were presented with packets with no disclosed salaries, they were more likely to report doing external research to help determine salary levels (82.0% vs. 73.5% for those who saw zero rather than four or eight disclosed salaries, one-sided p = 0.09).²¹ This concords with the findings of B arach and Horton (2021), which shows that when employers could not observe full compensation histories, they asked applicants more questions and spent more time acquiring additional information.

Nonetheless, they may have underestimated market wages for software engineers. To address this, Table G4 examines the subset of recruiters who receive packets of half-disclosing, half-silent candidates. These subjects address this question because the half of candidates who disclosed a number gave a reminder of general market wages to use as a benchmark for the silent candidates. However, in this sample, our results are very similar to the full sample—silent candidates are assumed to be adversely selected and to be similar to candidates who disclose around the 25th percentile of workers with the same observables. This implies our result is not likely an artifact of recruiter inexperience or lack of knowledge of market wages.

G Additional Empirical Analysis

²¹The question asked "How did you make judgments on the salary-related questions? Select all that apply," and the options were "Used my previous experience with salaries in this setting"; "Looked up salaries on a website like PayScale.com, Glassdoor.com, etc."; "Spoke with others familiar with salaries for software engineers"; and "Other." We considered the recruiter to "do research" if he or she reported looking up salaries or speaking with others.

	Female Candidate	Median Salary Cur- rent Empl (10K)	College Grad. Year	Disclosed Salary Cur- rent Empl (10K)	Amazon	Facebook	IBM	% of Sam- ple
All Candidates	0.50	9.97	2013.66	9.71	0.06	0.13	0.09	100.0
No Salary Prompt	0.50	9.97	2013.66	9.71	0.06	0.13	0.09	43.8
Has Salary Prompt	0.50	9.97	2013.66	9.71	0.06	0.13	0.09	56.3
No Disclosure	0.50	9.98	2013.67		0.06	0.13	0.10	40.6
Salary Disclosed	0.50	9.97	2013.65	9.71	0.06	0.12	0.09	59.4

Notes: This table shows the attributes of the fictitious job candidates overall, by whether their application included a salary history prompt, and by whether the candidate disclosed their salary in the application form. These are balanced by design.

	Female Re- cruiter	White	Black	3+ Yrs Exp	Hourly Rate	Asked Salary Input	% of Sam- ple
All Recruiters	0.75	0.52	0.23	0.71	44.07	0.98	100.0
No Salary Prompt	0.76	0.56	0.22	0.67	43.65	0.97	43.8
Has Salary Prompt	0.74	0.49	0.24	0.74	44.40	0.99	56.3
0 Salaries Disclosed	0.77	0.55	0.20	0.68	43.07	1.00	21.9
4 Salaries Disclosed	0.75	0.51	0.28	0.71	44.09	0.97	37.5
8 Salaries Disclosed	0.74	0.52	0.20	0.72	44.59	0.98	40.6
NoPrmpt 0Disc	0.72	0.56	0.19	0.66	44.97	1.00	12.5
NoPrmpt 4Disc MoreHigh	0.81	0.38	0.31	0.69	48.84	0.94	6.3
NoPrmpt 4Disc MoreLow	0.63	0.50	0.25	0.63	41.23	0.94	6.3
NoPrmpt 4Disc Mixed	0.88	0.69	0.31	0.69	37.32	1.00	6.3
NoPrmpt 8Disc AllHigh	0.75	0.63	0.13	0.63	41.88	0.88	3.1
NoPrmpt 8Disc AllLow	0.63	0.75	0.00	0.38	42.31	1.00	3.1
NoPrmpt 8Disc Mixed	0.88	0.56	0.25	0.88	46.13	1.00	6.3
Prmpt 0Disc	0.83	0.54	0.21	0.71	40.54	1.00	9.4
Prmpt 4Disc MoreHigh	0.81	0.56	0.13	0.75	42.88	1.00	6.3
Prmpt 4Disc MoreLow	0.63	0.50	0.44	0.75	46.38	0.94	6.3
Prmpt 4Disc Mixed	0.75	0.44	0.25	0.75	47.92	1.00	6.3
Prmpt 8Disc AllHigh	0.69	0.31	0.44	0.69	46.75	1.00	6.3
Prmpt 8Disc AllLow	0.56	0.50	0.06	0.75	40.06	0.94	6.3
Prmpt 8Disc Mixed	0.80	0.53	0.20	0.75	45.92	1.00	15.6

Table G2: Recruiter Balance

Notes: This table shows a subset of the demographics of our recruiting workforce of 256 recruiters, by whether they were shown applications with a salary history prompt or not, whether they saw 0, 4 or 8 candidates disclose a salary, and by combinations of prompt/no prompt, 0/4/8 salary disclosures, and distributions of amounts disclosed.

Panel A: TIOLI								
	Offer	Offer	Offer	Offer	Offer	Offer		
WTP	.93***		.77***	.91***		.77***		
	(.0087)		(.017)	(.0094)		(.018)		
Outside Offer,								
50th Percentile		.85***	.19***		.83***	.19***		
		(.018)	(.017)		(.021)	(.017)		
Candidate FEs				Y	Y	Y	_	
R^2	.92	.74	.94	.93	.74	.94		
Observations	2,048	2,048	2,048	2,048	2,048	2,048	_	
Panel B: Callback							_	
	Ca	llback	Callback	Callba	ck Ca	llback	Callback	Callback
WTP (10k)	.0	64***		.15***	• .0	72***		.14***
).)	0068)		(.013)).) (.	071)		(.013)
Outside Offer (10	lk),							
50th Percentile	•		.023***	1***	÷		.028***	09***
			(.0067)	(.011))		(.0068)	(.011)
Candidate FEs						Y	Y	Y
R^2		075	.009	.13		.13	.057	.17
Observations	2	,048	2,048	2,048	2	,048	2,048	2,048

Table G3: Correlates of Offers and Callback Decisions

Notes: This table shows the relationship between a selected group of the variables reported by our recruiters. Dependent variables are listed in the column header and all variables are explained in notes to Table 1. Standard errors are robust. + p < 0.10 * p < 0.05 ** p < 0.010 *** p < 0.001.

Table G4: Average Effect of Disclosing for Packets with Half of Salaries Disclosed

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	WTP	WTP	Outside Offer 50th %ile	Outside Offer 50th %ile	Offer	Offer	Outside Offer Range	Outside Offer Range
Salary Disclosed	0.50***	0.60**	0.61***	0.98***	0.57***	0.75***	-0.45	-0.87
	(0.13)	(0.20)	(0.12)	(0.17)	(0.11)	(0.17)	(0.33)	(0.63)
Female x Disclosed		-0.21		-0.74***		-0.36		0.83
		(0.26)		(0.21)		(0.23)		(0.63)
Female Disclosure Effect:								
Total		0.40		0.24		0.39		-0.04
p-value		0.01		0.11		0.01		0.70
Mean Non-Dislosers:								
All	10.30	10.30	9.48	9.48	9.65	9.65	2.85	2.85
Male	10.49	10.49	9.61	9.61	9.80	9.80	3.20	3.20
Female	10.11	10.11	9.35	9.35	9.51	9.51	2.50	2.50
R ²	0.08	0.07	0.11	0.12	0.11	0.11	0.00	0.00
Observations	768	768	768	768	768	768	768	768

Notes: This table shows estimates from versions of Equation 5 and mimics Table 2; the sample is restricted to data from recruiters who evaluated packets with exactly half of salaries disclosed (4 disclosed salaries, 4 non-disclosed salaries). All models include recruiter controls and candidate fixed effects. Dependent variables are listed in the column header and explained in notes to Table 1. Outcomes measured in dollars (e.g. WTP, Offer) are in \$10K increments. This table is the subset of the data presented in the main text. This table only shows data where a recruiter saw exactly half of the candidates disclosing their salary history versus all or none of the candidates disclosing as presented in Table 2. Robust standard errors are clustered at the recruiter level. + p < 0.10 * p < 0.05 * * p < 0.010 * * p < 0.001.

Table G5: Average Effect of Disclosing a High versus Low Salary forPackets with Half of Salaries Disclosed

	(1)	(2)	(3)	(4)
	WTP	Outside Offer 50th %ile	Offer	Outside Offer Range
Disclosed 25th %ile Salary	-0.09	-0.04	-0.05	-0.65+
	(0.15)	(0.14)	(0.13)	(0.38)
Disclosed 75th %ile Salary	1.13***	1.28***	1.21***	-0.24
-	(0.17)	(0.15)	(0.15)	(0.31)
Mean Non-Disclosers	10.34	9.64	9.72	2.65
R ²	0.22	0.30	0.25	-0.01
Observations	768	768	768	768

Notes: This table mimics Table 3 but the sample is restricted to data from recruiters who evaluated packets with exactly half of salaries disclosed (4 disclosed salaries, 4 non-disclosed salaries). All models include recruiter and spillover controls and both candidate and sub-treatment fixed effects as described in the text. Dependent variables are listed in the column header and explained in notes to Table 1. Outcomes measured in dollars (e.g. WTP, Offer) are in \$10K increments. Disclosed Xth %tile Salary means a candidate disclosed a salary at the Xth percentile within their specific firm. The omitted category is candidates who did not disclose a salary. Robust standard errors are clustered at the recruiter level. + p < 0.10 * p < 0.05 ** p < 0.010 *** p < 0.001

Table G6: Average Effect of Disclosing by Salary Amount for Packetswith Half of Salaries Disclosed

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	WTP	WTP	Outside Offer 50th %ile	Outside Offer 50th %ile	Offer	Offer	Outside Offer Range	Outside Offer Range
Salary Disclosed	-5.33***	-5.69***	-5.78***	-5.80***	-5.54***	-5.66***	0.32	0.80
-	(0.74)	(0.88)	(0.71)	(0.82)	(0.66)	(0.78)	(1.81)	(2.79)
Female x Disclosed		0.04		-0.15		-0.27		-1.44
		(0.89)		(0.83)		(0.78)		(2.37)
Female Disclosure Slope:								
Total		0.67		0.69		0.70		0.06
p-value		0.00		0.00		0.00		0.54
Mean Non-Dislosers:								
All	10.30	10.30	9.48	9.48	9.65	9.65	2.85	2.85
Male	10.49	10.49	9.61	9.61	9.80	9.80	3.20	3.20
Female	10.11	10.11	9.35	9.35	9.51	9.51	2.50	2.50
R ²	0.27	0.27	0.36	0.36	0.31	0.32	-0.01	-0.01
Observations	768	768	768	768	768	768	768	768

Notes: This table mimics Table 4 but the sample is restricted to data from recruiters who evaluated packets with exactly half of salaries disclosed (4 disclosed salaries, 4 non-disclosed salaries). All models include recruiter and spillover controls and both candidate and sub-treatment fixed effects as described in the text. This table shows estimates from versions of Equation 5 that include interactions with gender. Dependent variables are listed in the column header and explained in notes to Table 1. Outcomes measured in dollars (e.g. WTP, Offer) are in \$10K increments. Robust standard errors are clustered at the Recruiter level. + p < 0.10 * p < 0.05 ** p < 0.010 *** p < 0.001

	(1) Outside Offer 5th %ile	(2) Outside Offer 5th %ile	(3) Outside Offer 95th %ile	(4) Outside Offer 95th %ile	(5) 2 Other Offers	(6) 2 Other Offers
Salary Disclosed	1.17***	1.67***	0.25	0.42	-0.01	-0.03
	(0.14)	(0.16)	(0.27)	(0.43)	(0.03)	(0.04)
Female x Disclosed		-1.01***		-0.34		0.03
		(0.11)		(0.37)		(0.04)
Female Disclosure Effect:						
Total		0.67		0.08		0.01
p-value		0.00		0.65		0.82
, Mean Non-Dislosers:						
All	8.16	8.16	11.01	11.01	0.54	0.54
Male	8.23	8.23	11.43	11.43	0.57	0.57
Female	8.09	8.09	10.59	10.59	0.50	0.50
R ²	0.27	0.28	0.06	0.06	0.03	0.03
Observations	2048	2048	2048	2048	2048	2048
	(7)	(8)	(9)	(10)	(11)	(12)
	Candidate	Candidate	Firm	Firm	Both	Both
	Searches	Searches	Searches	Searches	Search	Search
Salary Disclosed	-0.02	0.00	0.02	-0.00	-0.01	-0.04
2	(0.03)	(0.03)	(0.03)	(0.03)	(0.03)	(0.03)
Female x Disclosed		-0.04		0.04		0.07 +
		(0.04)		(0.04)		(0.04)
Female Disclosure Effect:						
Total		-0.04		0.04		0.03
p-value		0.23		0.23		0.41
Mean Non-Dislosers:				0		
All	0.42	0.42	0.58	0.58	0.41	0.41
Male	0.39	0.39	0.61	0.61	0.45	0.45
Female	0.46	0.46	0.54	0.54	0.37	0.37
R ²	0.03	0.03	0.03	0.03	0.00	0.00
Observations	2048	2048	2048	2048	2048	2048

Table G7: Average Effect of Disclosing Salary for Additional Outcomes

Notes: This table mimics Table 2 but for additional outcomes we collected. All models include recruiter controls and candidate fixed effects. This table shows estimates from versions of Equation 5. Dependent variables are listed in the column header. Columns 5-12 are binary outcomes: 2 Other Offers means recruiter thinks the candidate will have 2 or more outside offers (as opposed to 1 or fewer); Candidate Searches means recruiter thinks outside offer likely comes from candidate aggressively pursuing outside options; Firm Searches means recruiter thinks outside offers likely come from other firms aggressively pursuing this candidate; and Both Search means recruiter believes outside offers offers one from both candidate and other firms pursuit. Outcomes measured in dollars (e.g. the outside offers) are in \$10K increments. Robust standard errors are clustered at the recruiter level. + p < 0.10 * p < 0.05 * p < 0.010 *** p < 0.001

comes							
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	Outside Offer	Outside Offer	Got Only	2 Other	Candidate	Firm	Both
	5th %ile	95th %ile	Offer	Offers	Searches	Searches	Search
+\$10k from Firm	0.73***	0.72***	-0.01	0.02	-0.01	0.01	0.01
	(0.06)	(0.21)	(0.01)	(0.02)	(0.02)	(0.02)	(0.02)
+\$10k from Male	0.62***	0.23	-0.02	-0.02	0.03	-0.03	-0.03
	(0.08)	(0.27)	(0.02)	(0.03)	(0.03)	(0.03)	(0.03)
+\$10k within Firm	0.71***	0.75***	0.01	0.01	-0.02	0.02	0.04*

(0.01)

0.50

0.14

0.13

0.01

2048

(0.02)

0.17

0.89

0.25

0.05

2048

(0.02)

0.09

0.82

0.07

0.06

2048

(0.02)

0.09

0.82

0.07

0.06

2048

(0.02)

0.14

0.14

0.02

0.03

2048

(0.06)

0.04

0.87

0.04

0.12

2048

(0.06)

0.04

0.63

0.14

0.54

2048

pF-M

p F-W

p M-W

Observations

R²

Table G8:	Effect of An	Extra Dollar	Decomposed	for	Additional	Out-
comes						

Notes: This table mimics Table 5 but for additional outcomes we collected. All models include recruiter and spillover controls and both candidate and sub-treatment fixed effects as described in the text. This table shows estimates from Equation 6 which decomposes additional dollars of salary disclosure into a firm-specific offset for the candidate's employer ("+\$10k from Firm," some firms pay higher or lower to everyone on average); a gender offset ("+\$10k from Male", which mimics real-world gender gaps); and from having a higher or lower salary within the current firm's distribution ("+\$10k within Firm", note this also is combined with some random noise that was included in the salaries.). Dependent variables are listed in the column header and explained in notes to Table G7. Outcomes measured in dollars (e.g. the outside offers) are in \$10K increments. *p*-values for comparisons of coefficients within the same model are provided in the 2nd panel, where for example *p* F-M is the p-value testing that the coefficient from "+\$1 from Firm" = the coefficient on "+\$1 from Male". Robust standard errors are clustered at the recruiter level. + p < 0.10 * p < 0.05 * p < 0.010 *** p < 0.001

Table G9: Average Effect of Disclosing by Salary Amount for Additional Outcomes

	(1) Outside Offer 5th %ile	(2) Outside Offer 5th %ile	(3) Outside Offer 95th %ile	(4) Outside Offer 95th %ile	(5) 2 Other Offers	(6) 2 Other Offers
Salary Disclosed	-6.23*** (0.54)	-6.24*** (0.60)	-6.42*** (0.98)	-7.37*** (1.74)	-0.02 (0.15)	-0.08 (0.18)
Female x Disclosed	(010 1)	-0.11 (0.37)	(000)	0.91 (1.57)	(0120)	0.05 (0.16)
Disclosed x Amt Disclosed	0.71*** (0.06)	0.71*** (0.06)	0.69*** (0.12)	0.74*** (0.19)	0.01 (0.01)	0.01 (0.02)
Female x Disclosed x Amt Disclosed	. ,	0.03 (0.04)		-0.01 (0.15)	. ,	0.00 (0.02)
Female Disclosure Effect:						
Total p-value		$\begin{array}{c} 0.74 \\ 0.00 \end{array}$		0.73 0.00		$\begin{array}{c} 0.01 \\ 0.40 \end{array}$
Mean Non-Dislosers: All	8.16	8.16	11.01	11.01	0.54	0.54
Male Female	8.23 8.09	8.23 8.09	11.43 10.59	11.43 10.59	0.57 0.50	0.57 0.50
R ² Observations	0.54 2048	0.54 2048	0.12 2048	0.12 2048	0.05 2048	0.05 2048
Observations	(7) Candidate Searches	(8) Candidate Searches	(9) Firm Searches	(10) Firm Searches	(11) Both Search	(12) Both Search
Salary Disclosed	0.05 (0.14)	0.10 (0.17)	-0.05 (0.14)	-0.10 (0.17)	-0.15 (0.15)	-0.12 (0.18)
Female x Disclosed		-0.01 (0.17)		0.01 (0.17)	× ,	-0.13 (0.17)
Disclosed x Amt Disclosed	-0.01 (0.01)	-0.01 (0.02)	0.01 (0.01)	0.01 (0.02)	0.02 (0.02)	0.01 (0.02)
Female x Disclosed x Amt Disclosed		-0.01 (0.02)		0.01 (0.02)	· ·	0.02 (0.02)
Female Amount Disclosed Slope:						
Total p-value Mean Non-Dislosers:		-0.02 0.26		0.02 0.26		0.03 0.04
All Male	0.42 0.39	0.42 0.39	0.58 0.61	0.58 0.61	$0.41 \\ 0.45$	$0.41 \\ 0.45$
Female R ²	0.46 0.06	0.46 0.06	0.54 0.06	0.54 0.06	0.37 0.03	0.37 0.03
Observations	2048	2048	2048	2048	2048	2048

Notes: This table mimics Table 4 but for additional outcomes we collected. All models include recruiter and spillover controls and both candidate and sub-treatment fixed effects as described in the text. This table shows estimates from versions of Equation 5 that include interactions with gender. Dependent variables are listed in the column header and explained in notes to Table G7. Salary Amounts and outcomes measured in dollars (e.g. the outside offers) are in \$10K increments. Robust standard errors are clustered at the Recruiter level. + p < 0.10 * p < 0.05 * p < 0.010 * ** p < 0.001

Table G10: Average Effect of Disclosing Salary for Prompted versus Unprompted Disclosure

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	WTP	WTP	Outside Offer	Outside Offer	Offer	Offer	Outside Offer	Outside Offer
	(10k)	(10k)	50th %ile	50th %ile	(10k)	(10k)	Range	Range
Salary Disclosed	0.60**	0.90***	0.82***	1.27***	0.64**	1.02***	-0.45*	-0.52*
	(0.21)	(0.24)	(0.19)	(0.22)	(0.21)	(0.23)	(0.21)	(0.22)
Salary Disclosed x Prompt	0.15	0.09	0.05	0.04	0.17	0.09	-0.88*	-1.44+
	(0.29)	(0.33)	(0.27)	(0.31)	(0.28)	(0.31)	(0.41)	(0.75)
Prompt on Application	-0.03	0.10	0.20	0.29	-0.01	0.07	0.66	1.23
	(0.26)	(0.28)	(0.25)	(0.26)	(0.25)	(0.27)	(0.41)	(0.76)
Female x Disclosed		-0.61**		-0.91***		-0.76***		0.15
		(0.19)		(0.17)		(0.17)		(0.11)
Female x Salary Disclosed x Prompt		0.12		0.03		0.16		1.13
		(0.27)		(0.23)		(0.25)		(0.78)
Female x Prompt on Application		-0.25		-0.18		-0.15		-1.14
		(0.21)		(0.18)		(0.19)		(0.77)
Mean Unprompted Non-Dislosers:								
All	10.27	10.27	9.35	9.35	9.61	9.61	2.57	2.57
Male	10.40	10.40	9.44	9.44	9.72	9.72	2.66	2.66
Female	10.14	10.14	9.26	9.26	9.50	9.50	2.49	2.49
R ²	0.18	0.18	0.25	0.26	0.20	0.21	0.02	0.02
Observations	2048	2048	2048	2048	2048	2048	2048	2048

Notes: This Table mimics Table 2 and adds in controls and interactions with whether the disclosure was in response to prompt or if it was unprompted. All models include recruiter controls and candidate fixed effects. This table shows estimates from versions of Equation 5. Dependent variables are listed in the column header and explained in notes to Table 1. Salary Amounts and outcomes measured in dollars (e.g. WTP, offer) are in \$10K increments. Robust standard errors are clustered at the recruiter level. + p < 0.10 * p < 0.05 * p < 0.010 * * p < 0.001

Table G11: Average Effect of Disclosing by Salary Amount for Promptedversus Unprompted Disclosure

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	WTP	WTP	Outside Offer	Outside Offer	Offer	Offer	Outside Offer	Outside Offer
	(10k)	(10k)	50th %ile	50th %ile	(10k)	(10k)	Range	Range
Salary Disclosed	-6.21***	-6.20***	-7.18***	-7.70***	-6.65***	-6.50***	0.04	-0.96
	(0.63)	(0.73)	(0.58)	(0.64)	(0.58)	(0.66)	(0.90)	(1.78)
Disclosed x Amt Disclosed (10K)	0.70***	0.71***	0.81***	0.86***	0.75***	0.75***	-0.01	0.07
	(0.06)	(0.07)	(0.06)	(0.06)	(0.06)	(0.06)	(0.11)	(0.16)
Salary Disclosed x Prompt	-0.40	-1.23+	-0.63	-0.47	-0.02	-0.90	-0.97+	-1.24
	(0.58)	(0.68)	(0.49)	(0.59)	(0.53)	(0.63)	(0.53)	(1.02)
Disclosed x Amt Disclosed (10k) x Prompt	0.06	0.07	0.07	0.02	0.02	0.05	0.01	-0.02
	(0.05)	(0.06)	(0.05)	(0.06)	(0.05)	(0.06)	(0.04)	(0.05)
Prompt on Application	-0.15	0.00	0.07	0.00	-0.14	0.00	0.65	0.00
	(0.29)	(.)	(0.27)	(.)	(0.28)	(.)	(0.43)	(.)
Female x Disclosed		0.28		0.79		-0.04		0.73
		(0.67)		(0.57)		(0.61)		(1.43)
Female x Disclosed x Amt Disclosed (10K)		-0.00		-0.06		0.02		-0.05
		(0.07)		(0.06)		(0.06)		(0.13)
Female x Salary Disclosed x Prompt		0.04		-0.80		0.31		0.70
		(0.76)		(0.63)		(0.67)		(0.91)
Female x Prompt on Application		-0.27		-0.18		-0.16		-1.14
		(0.21)		(0.17)		(0.19)		(0.78)
Female x Disclosed x Amt Disclosed (10k) x Prompt		0.03		0.10		-0.00		0.04
		(0.07)		(0.06)		(0.07)		(0.06)
Mean Unprompted Non-Dislosers:								
All	10.27	10.27	9.35	9.35	9.61	9.61	2.57	2.57
Male	10.40	10.40	9.44	9.44	9.72	9.72	2.66	2.66
Female	10.14	10.14	9.26	9.26	9.50	9.50	2.49	2.49
R ²	0.36	0.37	0.52	0.52	0.42	0.43	0.02	0.02
Observations	2048	2048	2048	2048	2048	2048	2048	2048

Notes: This Table mimics Table 4 and adds in controls and interactions with whether the disclosure was in response to prompt or if it was unprompted. All models include recruiter and spillover controls and both candidate and sub-treatment fixed effects. This table shows estimates from versions of Equation 5. Dependent variables are listed in the column header. Salary Amounts and outcomes measured in dollars (e.g. WTP, offer) are in \$10K increments. Robust standard errors are clustered at the recruiter level. + p < 0.10 * p < 0.05 * * p < 0.010 * ** p < 0.00

Table G12: Who are Complier Types?

Panel A: Always Disclosers

Panel A: Always Disclosers	Always	Always	Always	Always	Always	Always	Always	Always
Female	Discloser 12***	Discloser	Discloser	Discloser	Discloser	Discloser 11***	Discloser 11***	Discloser 13***
High Salary w/in Firm	(.028)	017				(.028) 016	(.03) 012	(.045) 035
Salary (Normalized)		(.028)	.054***			(.028) .048***	(.028) .046***	(.043) .067***
Occupation's Average Salary (Norm)			(.016)	.031**		(.012)	(.011)	(.023)
Industry's Average Salary (Norm)				(.015)	.021			
					(.015)			.038
Female× High Salary w/in Firm								(.059)
Female× Salary (Norm)								.053 (.087)
High Salary w/in Firm× Salary (Norm)								0003 (.068)
Fem.× High Salary @Firm× Salary (Norm)								086 (.11)
Industry FEs Occupation FEs							Y Y	Y Y
R ² Observations	.018 1,006	.00038 1,006	.013 1,006	.0049 1,006	.0022 1,006	.029 1,006	.059 1,005	.061 1,005
Observations	1,000	1,000	1,000	1,000	1,000	1,000	1,005	1,005
anel B: Ban Compliers								
	Ban Complier	Ban Complier	Ban Complier	Ban Complier	Ban Complier	Ban Complier	Ban Complier	Ban Complie
Female	.11*** (.031)	· ·			*	.1*** (.031)	.12*** (.034)	.11** (.05)
High Salary w/in Firm	(1001)	034 (.032)				035 (.031)	043 (.032)	041 (.046)
Salary (Normalized)		(.032)	045***			04***	044***	076**
Occupation's Average Salary (Norm)			(.016)	0074		(.013)	(.015)	(.028)
Industry's Average Salary (Norm)				(.016)	.00013			
Female× High Salary w/in Firm					(.016)			.008
Female× Salary (Norm)								(.066) 079
High Salary w/in Firm× Salary (Norm)								(.1) .019
Fem.× High Salary @Firm× Salary (Norm)								(.071) .11
Industry FEs							Y	(.12) Y
Occupation FEs R ²	.012	.0011	.0076	.00022	7.1e-08	.019	Y .047	Y .05
Observations	1,006	1,006	1,006	1,006	1,006	1,006	1,005	1,005
Panel C: Never Disclosers	Never	Never	Never	Never	Never	Never	Never	Never
Female	Discloser .014	Discloser	Discloser	Discloser	Discloser	Discloser .012	Discloser 0018	Discloser .023
High Salary w/in Firm	(.025)	.048*				(.025) .048*	(.027) .051**	(.038) .07**
Salary (Normalized)		(.025)	0088			(.025) 0082	(.025) 0026	(.035) .0093
Occupation's Average Salary (Norm)			(.0066)	021*		(.0066)	(.0074)	(.02)
Industry's Average Salary (Norm)				(.012)	018			
Female× High Salary w/in Firm					(.012)			042
Female× Salary (Norm)								(.051) .023
High Salary w/in Firm× Salary (Norm)								(.081) 031
Fem.× High Salary @Firm× Salary (Norm)								(.06) 0068
Industry FEs							Y	(.1) Y
Occupation FEs R ²	.00033	.0037	.00053	.0028	.0022	.0045	Y .024	Y .025
R Observations	1,006	1,006	1,006	1,0028	1,0022	1,0045	1,005	1,005

Notes: + p < 0.10 * p < 0.05 ** p < 0.010 *** p < 0.001

Table G13: Full vs. Partial Ban: Effect on Expected Salary Offers (Salary Offer \times Callback)

Salary × Callbacks (Full Ban)

 $Salary \times Callbacks$ (Partial Ban)

	Women	Men	Ratio
D			
No Ban	64919.37	70646.28	0.92
	(2391.37)	(2717.71)	(0.04)
Ban	63937.62	64882.21	0.99
	(2886.18)	(2487.74)	(0.05)
Ban-No Ban	-981.75	-5764.07	0.07
	(3748.16)	(3684.40)	(0.07)
<i>p</i> -value	0.79	0.12	0.31

Notes: This table shows the effect of a salary history ban on the salary offer multiplied by a binary variable for if the candidate was recommended for a callback. The left panel shows this for a full ban and the right panel is for a partial ban. These are the effects for a "Full Ban" meaning a ban where salary history may not be asked at any stage in the hiring process. The effects for a "Parial Ban" meaning a ban of prompting job candidates to disclose on the job application, but being able to seek salary information at a later stage in the hiring process. Standard errors are robust.

H Experimental Materials

H.1 Sample Job Application: Salary History Asked + Candidate Discloses

Samantha Application Details for Software Engineering position All sections are required except where noted. For candidates who are interviewed, all information entered below will be verified. **Candidate Information** Candidate Id: 774 Mailing Address: City/State: ZIP: URL: http:// Phone: (Email: Are you legally authorized to work in the US? Y Are you over the age of 18?: Y Are you willing to relocate for this position? Y Will you now (or in the future) require visa sponsorship? N **Employment History (Last Three Jobs)** Title: Software Engineer Company Name: IBM Location: San Jose, CA Dates: 01/2015 - Present Position Description, Duties, Responsibilities: * Developing and implementing new feedback system for user concerns, bugs, and defect tracking regarding use and functionality of new interfaces * Coding web designed interfaces using Java, XML, XSL, AJAX, and JWS. * Implement the command-line interface for the Universal Authentication Protocol (UAP) in E-directory. Title: Software Developer Company Name: Amazon Location: Seattle, WA Dates: 05/2014 - 01/2015 Position Description, Duties, Responsibilities: * Developed code and unit tests in Python for server-side and in JavaScript for web components. * Deployed and tested code on Linux-based EC2 instances in a distributed AWS cloud environment. * Created and maintained automated jobs to build and test software. * Developed and implemented working plans for the formulation of front and back-end web applications. * Developed various algorithms to mitigate program interference. Title: Programming Intern Company Name: Intraix Location: Aver Rajah Crescent, SG Dates: 05/2013 - 08/2013 Position Description, Duties, Responsibilities: Automated black box and white box tests for an Android application "Klug," using Appium and Espresso framework. This helped developers expand features without much worry of breaking current functionalities. Salary History Annual Base Salary at Current or Most Recent Job: \$96,000 Education Institution: Georgia Institute of Technology Dates: 2010 - 2014 Graduated? Y Location: Atlanta, GA Level: BS (Bachelor of Science) Subject/Major: Computer Science Relevant Coursework Database and Information Management Systems, Java, Analysis of Algorithms, Data Sytems, Matlab for Programmers, and Compiler Design

Additional Skills and Information

Experience developing in Java, HTML/CSS, JavaScript, Node.js, Ruby, Ruby on Rails, Shell, Python, SQL, LATEX.

H.2 Sample Job Application: Salary History Asked + Candidate Does Not Disclose

Christopher

Application Details for Software Engineering position

All sections are required except where noted. For candidates who are interviewed, all information entered below will be verified.

Candidate Information

Candidate Id: 721 Mailing Address:		City/State:		
ZIP: Phone: (- Email:	URL: http://		
Are you legally authorized to work in the US? ${\sf Y}$		Are you over the age of 18?: Y		
Are you willing to relocate for this position? ${\sf Y}$		Will you now (or in the future) require visa sponsorship? N		

Employment History (Last Three Jobs)

Title: Programmer Company Name: Apple Location: Cupertino, CA

Position Description, Duties, Responsibilities:

Research, design, and implement scalable applications for information identification, extraction, analysis, retrieval, and indexing. Direct software design and development while remaining focused on client needs. Collaborate closely with other team members to plan, design, and develop robust solutions. Maintain front-end admin interface as well as back data processing.

Title: Programmer Company Name: Verizon Communications, Inc. Location: New York, NY Dates: 07/2014 - 10/2015

Position Description, Duties, Responsibilities:

Designed, developed, and integrated software with test systems hardware for test engineering applications. Supported the design and testing of space systems software in all program phases, from initial design through coding, testing, and integration. Member of team responsible for developing a new high-end software package. Led team of 3 engineers to manage Windows client (C++) including feature development, debugging, and update release.

Title: Summer Programming Associate Company Name: Facebook Location: Menlo Park, CA Dates: 06/2013 - 08/2013

Position Description, Duties, Responsibilities:

Intern on the Sales Platform team within Core Ads, which deals primarily with making tools to help salespeople make sales, usually by connecting them to advertisers. Worked on improving the infrastructure and data quality of our platform that helps sales teams find their clients. Languages/technologies: Hack (PHP), Python, Dataswarm.

Salary History (optional)

Annual Base Salary at Current or Most Recent Job:

Education

Institution: California Institute of Technology

Location: Pasadena, CA Dates: 2010 - 2014 Subject/Major: Computer Science

14 Graduated? Y

Dates: 10/2015 - Present

Level: BS (Bachelor of Science) Relevant Coursework:

Artificial language, hardware systems, analysis of algorithms. programming abstractions, data structures and algorithms

Additional Skills and Information

Production code launched using C/C++, Java, Javascript, Python, Perl. Back-end and research experience using Linux shell scripting, R, PiCloud/Multivac, Sawzall, MapReduce.

1

H.3 Sample Job Application: Salary History Not Asked + Candidate Does Not Disclose

Sarah Application Details for Software Engineering position All sections are required except where noted. For candidates who are interviewed, all information entered below will be verified. **Candidate Information** Candidate Id: 1724 Mailing Address: City/State: ZIP: Phone: (Email: URL: http:// Are you legally authorized to work in the US? Y Are you over the age of 18?: Y Are you willing to relocate for this position? Y Will you now (or in the future) require visa sponsorship? N Employment History (Last Three Jobs) Title: Coder Company Name: Facebook Location: Menlo Park, CA Dates: 06/2014 - Present Position Description, Duties, Responsibilities: Enhancing existing web applications to meet current standards. Constructing complex queries using SQL in the IBM DB2 Database. Designing technical structure and modules for a new and better UX. Collaborating with senior developers to execute client work. Introducing automated acceptance and unit tests, while increasing coverage. Title: Software Architect Company Name: Dell Location: Round Rock, TX Dates: 06/2013 - 06/2014 Position Description, Duties, Responsibilities: Participate in application modification and development of new applications to meet business needs. Provide full life-cycle project expertise. Project work focused on business applications and e-business solutions. Responsibilities included application integration and development using .NET including C#, ASP.Net, WinForms, MS Exchange, and Microsoft Sharepoint Portal Server. Title: Summer Coding Fellowship Company Name: Apple Location: Cupertino, CA Dates: 05/2012 - 08/2012 Position Description, Duties, Responsibilities: Built an automated framework on the Apple Maps Team for validating the internal pipeline that manages how different layers of maps data integrate using Python. Education Institution: Cornell University Dates: 2009 - 2013 Graduated? Y Location: Ithaca, NY Level: BS (Bachelor of Science) Subject/Major: Computer Science **Relevant Coursework:**

Systems Programming and Machine Organization, Privacy and Technology, Data Science I, Networks, Computing Hardware, Cloud Computing.

Additional Skills and Information

Skills: JS, Java, XPages, Flex / AIR, Processing, Git, Eclipse, HTML.

1

H.4 Sample Job Application: Salary History Not Asked + Candidate Volunteers

Tyler

Application Details for Software Engineering position

All sections are required except where noted. For candidates who are interviewed, all information entered below will be verified.

Candidate Information

Candidate Id: 62	andidate Id: 621 Mailing Address:			City/State:
ZIP: Ph	none: (- Email:		URL: http://
Are you legally authorized to work in the US? ${\sf Y}$		Are you over the age of 18?: Y		
Are you willing to relocate for this position? Y		Will you now (or in the future) require visa sponsorship? N		

Employment History (Last Three Jobs)

Title: Developer Company Name: Amazon Location: Seattle, WA Dates: 02/2014 - Present

Position Description, Duties, Responsibilities:

- Develop automated REST API test cases to ensure proper error handling.

- Conduct regression tests on internal and external products and services in order to successfully integrate new solutions to existing systems.

- Review and approve code releases from development and marketing departments. ensure thorough client policy compliance.

Title: Coder Company Name: Google Location: Mountain View, CA

Position Description, Duties, Responsibilities:

- Researched emerging technologies for database and network storage solutions by reviewing case studies and functionality to determine low-cost, but effective, models for supported environments.

- Provided leadership and decision making to impact infrastructure changes that included upgrading the Oracle database schema, applying new versions of Dart Enterprise, and implementing a virtualized hardware environment to reduce footprint and minimize data center presence.

Title: Software Development Trainee Company Name: GE Healthcare Location: Little Chalfont, UK Dates: 05/2012 - 08/2012

Position Description, Duties, Responsibilities:

Reduced waiting time to pull information from multiple systems - requests that used to take days, now only take minutes. Also worked closely with other IT professionals to design, test, and implement APIs in support of major ERP systems.

Education

Institution: University of Illinois at Urbana-Champagne Location: Champaign, IL Dates: 2009 - 2013 Graduated? Y

Level: BS (Bachelor of Science)

Subject/Major: Computer Science

Dates: 05/2013 - 02/2014

Relevant Coursework:

C++, Java, Microprocessor systems, Cryptography, Human-computer interface technology, Computer networks, and Large scale systems

Additional Skills and Information

Skilled in Python (Django), Java, Ruby on Rails, JavaScript (AngularJS, jQuery), SQL, PHP, HTML, CSS. I make about \$125,000 per year right now (pre-bonus).

1

H.5 Recruiter Instructions

Instructions

Thank you for your help screening our candidates. Please read these instructions carefully and completely before you begin this task.

1 About our Hiring Needs

We are interested in finding candidates for a full-stack software engineering position at a mid-sized software start-up company. Qualified candidates should have a working understanding of hardware systems infrastructure, creating and manipulating databases, writing back-end code in one or more languages (e.g., Ruby, Java, Python, C#), and writing front-end code in one or more languages (e.g., HTML, Javascript). Other responsibilities may include project management and technical documentation. Our company has locations in several cities throughout the United States and many of our software engineers work remotely; location will be determined in consultation with the candidate after an offer has been made.

Additional details about our opening are available in section 5.B.

2 Your Task

We will provide you with candidates' responses to our online job application form. We ask that you review this information and answer a few questions. In particular, we will ask you about:

- Whether we should interview the candidate
- What salary we should offer or accept if they pass our interview
- Additional questions about potential salary ranges

At this stage, we are interested in identifying worthy candidates. In that sense, we do not have a fixed number of positions so you should let us know about any candidate in our applicant pool that would be a good match for this position.

Software engineers currently at our firm make between \$70,000 and \$120,000. You should not feel constrained by our current range, and we welcome your own research about what candidates should be paid. We also offer benefits including health insurance, stock and a performance-based annual bonus. However, our questions for you today will be about the cash component (annual base salary) of compensation only.

2 Compensation for you, our recruiter, for this task

For your assistance with this task, you will be paid hourly (with a maximum of 2 hours allowable), plus a bonus. You can read the details about the bonus calculation in the appendix to these instructions, but we'll summarize it here:

1. We care about spending recruiting energy on candidates we're likely to hire -- candidates who will impress us in interviews and will accept our offers.

2. We care about the difference between what we pay candidates and the value they bring to our company. It's worth paying more for salaries, but only if they bring more value (and/or if they're more likely to accept). We want your decisions to consider value, cost, and probability of acceptance.

We will interview all candidates you suggest. We may also interview candidates you did not suggest upon recommendation from others at our company.

Please note, we do not negotiate salaries with candidates.

3 Your Feedback about the Candidates

We will provide an online form for you to fill in your evaluations to make it easier to work together without too much back-and-forth. There will be six sets of questions about the candidates themselves, and a few quick questions about yourself.

4 Additional Information

Ultimately our staff are very busy and not available to answer questions as you review these applications.

Please do not contact any of these candidates. We are asking you only to evaluate them and send us your private assessments. Someone from our staff will take the next step with the candidates. To prohibit you from contacting them, we have blacked out their contact information in the attached application forms.

Our hiring philosophy is to make interview decisions based on what is submitted. Therefore, please do not consult any information on individual candidates outside of the packets we send you. For example, do not look up the candidates on Google or LinkedIn.

5 Appendix

The remainder of this document includes:

- Some additional details about your bonus payment.
- Additional information about the job requirements for full-stack software engineer.

5.A Exact Formula for Calculating Your Bonus

We will calculate a bonus associated with each candidate you review according to the guidelines below, and then sum them up across all candidates and pay you the full sum in addition to your hourly rate. The bonuses will be paid after we have completed our interview and hiring decisions - approximately 45 days (or sooner) after your complete the task.

For candidates who are hired, we will examine their performance and trajectory about four weeks after the candidate starts work. We'll rate the newly hired candidate on three dimensions using the one-through-three scale outlined on the next page.

We will add up the candidate's three scores, for a total score ranging between 3 and 9. We then multiply that total score by five, and subtract [the candidate's salary / 100,000]. This is your bonus for each newly hired candidate.

Hired Candidate Bonus= 5 (Technical Score + Innova ion Score + Leadership Score) - Salary/100,000

As you know, you'll help set our workers' salaries through your feedback in this task.¹ This bonus gives you the incentive to find candidates who deliver a lot of value to our company above the salary we need to pay them.

For candidates who are NOT hired -- either because we don't make them an offer, or because they reject our offer -- your Hired Candidate Bonus for that candidate will be zero.

We will also subtract \$5 from your overall bonus for everyone you suggest interviewing who isn't hired. This is to encourage you to be a little bit selective about forwarding candidates who have a realistic shot at joining our company. If we hire someone who you didn't suggest interviewing, we'll calculate the Hired Candidate Bonus as if you suggested interviewing that candidate. Also: If you suggest interviewing a candidate and the candidate declines to be interviewed, we would count this as a failed interview.

Please note: It would (in theory) be possible to earn a negative overall bonus. If this happens, we will set the overall bonus to \$0.

¹ In one of our questions for you, we'll ask you what we should offer the candidate as a take-it-or-leave-it offer. For candidates who accept, we'll use that salary in the bonus calculation. We'll also ask you what to do if a candidate instead approaches us with a take-it-or-leave-it offer. If you guide us to accept those offers in some circumstances, then we'll use those salaries in the formula above.

Evaluation Dimensions

- A. Technical Score
- B. Innovation Score
- C. Leadership Score

Examples of Performance in Each Dimension

A. Technical Score:

Rating 1 (Low): Gaining command of all core technologies and practices used in our firm's engineering team. Able to begin developing and productionizing low to moderate complexity modules.

Rating 2 (Middle): Reasonable command of core engineering systems. Shows comfort with owning reasonably high complexity modules.

Rating 3 (High): Responsible for driving, technically designing, implementing and productionizing high impact projects with the help of teams if needed. Can own and deliver on very large mission-critical projects that impact the company in a verifiable way.

B. Innovation Score:

Rating 1 (Low): Responsible for implementing specifications developed by senior engineers and product managers. Does not develop products.

Rating 2 (Middle): Develops incrementally innovative ideas that can be successfully patented. Does not take leadership of developing new products, features and lines of business.

Rating 3 (High): Develops patentable ideas that lead to breakthrough improvements. Comes up with ideas to expand their projects and may also have a reasonable free-hand in developing and executing on them.

C. Leadership Score:

Rating 1 (Low): Tech, design or architectural lead of a small team/project, but could not have direct reports.

Rating 2 (Middle): Be able to mentor engineers in the team, giving technical guidance, code reviews, and ultimately be able to take responsibility of delivering small projects end-to-end on production.

Rating 3 (High): Leads complex initiatives and technically drives teams towards implementing and productionizing them. Promotes professional growth and development inside and outside the team. Actively takes steps to increase technical excellence across the organization.

5.B Additional Information about Job Opening for a Full-Stack Software Engineer

The position of software engineer will involve work on a specific project critical to a start-up's needs with opportunities to change projects and teams as the software engineer grows. Engineers are required to be multifaceted, display successful leadership abilities, and be enthusiastic to tackle new and challenging problems.

Responsibilities may include:

- Design, develop, test, deploy, maintain, and improve software
- Manage individual project priorities, deliverables, and deadlines
- Collaborate with other specialists in development teams
- Analyze and improve efficiency, scalability, and stability of various system resources

Minimum Qualifications:

- BA or BS degree in Computer Science or related technical field
- Experience with one or more general purpose programming languages including but not limited to: Java, C/C++, C#, Objective C, Python, JavaScript, or Go
- Experience working with two or more from the following: web application development, Unix/Linux environments, mobile application development, distributed and parallel systems, machine learning, information retrieval, natural language processing, networking, developing large software systems, and/or security software development
- Working proficiency and communication skills in verbal and written English

H.6 Recruiter Online Evaluation Form

- → C 습 🔒	https://tufts.ca1.qualtrics.com/jfe Q 🏫 🥕 🗿 😲 🐻 🚺 🝈 🗄
Close Preview	→ Restart Survey 🗘 - Draft - □ Place Bookmark - □
	Survey Completion 100%
Brief Feedback	on 8 Job Applications
candidates, so y	ask to take a maximum of two hours. We will ask you six sets of questions about the ou may want to keep any research open until the end of those questions. We will then uick background questions about yourself.
before hitting "S	his form will save your responses as you enter them, you can come back to a page ubmit" and complete the evaluations of candidates in any order. However, once you " on any particular page, you will not be able to revise your responses any more.
Click "Submit" to	get started.
	Submit

Please review the candidate information in the packets we sent you. In the table below, tell us:

1. Which candidates do you suggest interviewing? Then, for each candidate, assume he/she was interviewed and passed.

 Suggest a take-it-or-leave-it annual base salary offer for each candidate.
 Remember that it's worth paying more to make sure that higher quality candidates say yes. But we also are on a budget and do not want to overpay.

Please provide a salary offer even for candidates you do NOT suggest interviewing. Someone else at our company might suggest interviewing these candidates. If they pass, we will use your input for what salary to offer them.

For candidates you do NOT suggest interviewing, please enter the amount you think they should be offered were they to pass an interview - this may be helpful for us in the future. You may enter \$0, but only if that is what you truly intend, otherwise please enter a non-zero value.

	Candidate's	Candidate's	Interview	Your take-it-or- leave-it offer	Notes/Comments
	First name	ID #	Yes No	Amount In Dollars	(Optional)
1			0 0		
2			0 0		
3			0 0		
4			0 0		
5			0 0		
6			0 0		
7			0 0		
8			0 0		1

We will interview all candidates you suggest. We may also interview additional candidates. Interviewers will not know who suggested each candidate. We will consider your salary responses only after we finish interviewing and deciding which candidates to pursue.

If you could make a single take it or leave it offer, which candidate would you want to make it to?



Notes/Comments (Optional)

As a reminder, here are your answers to the previous questions.

First Name	ID	Interview	Salary
Andrew	1111	Yes	150,000
Christoper	2222	No	150,000
Emily	3333	Yes	150,000
Jessica	4444	No	150,000
Joshua	5555	Yes	150,000
Samantha	6666	No	150,000
Sarah	7777	Yes	150,000
Tyler	8888	No	150,000

Occasionally, candidates make us a take-it-of-leave-it offer (rather than us making one to them). Our policy is to take these proposals seriously. We will either accept or reject the candidate's offer, and the outcome is then final. We do not make counter offers or consider the candidate's follow-up offers. Candidates know this is our policy.

Please note: We do not feel that candidates who make us an offer are necessarily better or worse than those who wait for the offer.

We need your suggestions about how to respond in these settings. Please select the maximum offer we should accept for each candidate in this situation. For your reference, a table of the salary offers you suggested earlier can be found at the bottom of this page.

	Candidate's	Candidate's	Maximum take-it- or-leave-it offer we should accept	Notes/Comments
	First name	ID #		(Optional)
1	Andrew	1111	•	
2	Christoper	2222	T	
3	Emily	3333	T	
4	Jessica	4444	•	
5	Joshua	5555	T	
6	Samantha	6666	T	
7	Sarah	7777	T	
8	Tyler	8888	T	

First Name	ID	Interview	Salary
Andrew	1111	Yes	150,000
Christoper	2222	No	150,000
Emily	3333	Yes	150,000
Jessica	4444	No	150,000
Joshua	5555	Yes	150,000
Samantha	6666	No	150,000
Sarah	7777	Yes	150,000
Tyler	8888	No	150,000

As a reminder, here are your answers to some previous questions.

While we will interview all candidates you suggested, we may also interview additional candidates. For your responses, assume all candidates are interviewed and have passed. We will consider these responses only after we finish interviewing and deciding which candidates to pursue.

Suppose we were interested in hiring each of the candidates.

Tell us:

A salary so low, they'd be only 5% likely to take it?

A salary they'd be just as happy taking or rejecting (50% likely to accept or reject). A salary so high, we think they're very likely to take it? (95% likely to say yes)?

We are looking for 3 distinct salary values for each candidate. For example, for an entry level cashier job potential answers might be: there is a 5% chance they would take \$7.25, there is a 50% chance they would take \$9, and a 95% chance they would take \$15. Use your best judgement.

	Candidate's	Candidate's		Salary		Notes/Comments
	First name	ID #	5% likely to yes	50% likely to say yes	95% likely to say yes	(Optional)
1	Andrew	1111				
2	Christoper	2222				
3	Emily	3333				/
4	Jessica	4444				/
5	Joshua	5555				/
6	Samantha	6666				/
7	Sarah	7777				/
8	Tyler	8888				

While we will interview all candidates you suggested, we may also interview additional candidates. For your responses, assume all candidates are interviewed and have passed. We will consider these responses only after we finish interviewing and deciding which candidates to pursue.

We're interested in your opinion of each candidate's competing offers - please enter your assessments of their potential competing offers below

	Candidate's	Candidate's	Number of Competing Offers (not including current job)	Reason for Competing Offers	Notes/Comments
	First name	ID #			(Optional)
1	Andrew	1111	•	Y	
2	Christoper	2222	T	T T	1.
3	Emily	3333	•	×	
4	Jessica	4444	•	×	
5	Joshua	5555	•	×	
6	Samantha	6666	T	τ	
7	Sarah	7777	•	¥	
8	Tyler	8888	•	×	

Given your experience with this type of work, do you think that a job candidate like the ones you reviewed here would mis-report his/her most recent salary (even if we could verify past salary later)?

Yes, he/she may mis-report

No, he/she would not mis-report

Submit

Brief Feedback on 8 Job Applications

You said a job candidate might mis-report his/her most recent salary, if a candidate stated his/her current salary was \$90,000, what do you think this candidates true most current salary is?

Thank you for your work. We just have a few very short additional questions we would like you to answer

Upwork Profile URL:

How long have you been doing this type of work?

Less than 3 months

3 months to 1 year

1 to 3 years

3 to 10 years

Over 10 years

How often do you provide salary input during hiring?

Never

Sometimes

About half the time

Most of the time

Always

How did you make judgments on the salary related questions? Please check all that apply.

Used my previous experience with salaries in this setting

Looked up salaries on a website like payscale.com, glassdoor.com or others

Spoke with others who are familiar with salaries for software engineers

Other

The questions below are optional questions on background for statistical purposes.

I identify my gender as:		
Male		
Female		
Other		

I identify my ethnicity as (select as many as apply)

Black/African Caucasian Hispanic/Latin American
Hispanic/Latin American
Native American
Pacific Islander
Other

Do you have any suggestions to improvements in how we solicit advice from recruiters? This may include technical issues you had, or bigger picture things like other questions we should ask about evaluating candidates?



We thank you for your time spent taking this survey. Your response has been recorded.

References

- Acemoglu, Daron and David Autor, "Skills, tasks and technologies: Implications for employment and earnings," in "Handbook of labor economics," Vol. 4, Elsevier, 2011, pp. 1043–1171.
- Acquisti, Alessandro and Christina M Fong, "An experiment in hiring discrimination via online social networks," 2015.
- Adler, Laura, "What's a Job Candidate Worth? Explaining Variation in Pay-Setting Practices," *Working Paper*, 2020.
- **Barach, Moshe A and John J Horton**, "How do employers use compensation history? Evidence from a field experiment," *Journal of Labor Economics*, 2021, *39* (1), 193–218.
- Becker, Gary S, The economics of discrimination, University of Chicago press, 1957.
- Blau, Francine D, Peter Brummund, and Albert Yung-Hsu Liu, "Trends in occupational segregation by gender 1970–2009: Adjusting for the impact of changes in the occupational coding system," *Demography*, 2013, *50* (2), 471–492.
- **Dee, Thomas S**, "A teacher like me: Does race, ethnicity, or gender matter?," *American Economic Review*, 2005, 95 (2), 158–165.
- Giuliano, Laura, David I Levine, and Jonathan Leonard, "Manager race and the race of new hires," *Journal of Labor Economics*, 2009, 27 (4), 589–631.
- Glassdoor, "Progress on the Gender Pay Gap: 2019," White Paper, 2019.
- Goldin, Claudia, Sari Pekkala Kerr, Claudia Olivetti, and Erling Barth, "The expanding gender earnings gap: Evidence from the LEHD-2000 Census," *American Economic Review*, 2017, 107 (5), 110–14.
- Krueger, Alan B, "How computers have changed the wage structure: evidence from microdata, 1984–1989," *The Quarterly Journal of Economics*, 1993, *108* (1), 33–60.
- Mason, Malia F, Alice J Lee, Elizabeth A Wiley, and Daniel R Ames, "Precise offers are potent anchors: Conciliatory counteroffers and attributions of knowledge in negotiations," *Journal of Experimental Social Psychology*, 2013, 49 (4), 759–763.