

Original Research Article

The *COMT p.Val158Met* Polymorphism and Cognitive Performance in Adult Development, Healthy Aging and Mild Cognitive Impairment

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Key Words

Catechol-O-methyltransferase · Cognitive capacity · Cognitive flexibility · Aging · Mild cognitive impairment · Dopamine

Abstract

Background: The impact of genetic polymorphisms on cognition is assumed to increase with age as losses of brain resources have to be compensated for. We investigate the relation of catechol-O-methyltransferase (*COMT*) *p.Val158Met* polymorphism and cognitive capacity in the course of adult development, healthy aging and the development of mild cognitive impairment (MCI) in two birth cohorts of subjects born between 1930 and 1932 or between 1950 and 1952. **Methods:** Thorough neuropsychological assessment was conducted in a total of 587 participants across three examination waves between 1993 and 2008. The *COMT* genotype was determined as a restriction fragment length polymorphism after PCR amplification and digestion with *Nla*III. **Results:** Significant effects of the *COMT p.Val158Met* polymorphism were identified for attention and cognitive flexibility in the younger but not the older cohort. **Conclusion:** These results confirm the importance of the *COMT p.Val158Met* genotype on tasks assessing attention and cognitive flexibility in midlife but not in healthy aging and the development of MCI. Our findings suggest that the influence of *COMT* changes as a function of age, decreasing from midlife to aging.

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Introduction

Cognitive functioning is commonly linked to dopaminergic activity, and studies have associated age-related losses of dopaminergic activity with age-related cognitive decline [1]. Catechol-O-methyltransferase (*COMT*) is involved in the modulation of dopamine in the prefrontal cortex (PFC), where its resultant enzyme operates postsynaptically by inactivating neurotransmission [2, 3]. *COMT* contains a functional polymorphism, which enables the substitution of valine (*Val*) with methionine (*Met*). The *Met* allele results in the production of an enzyme that is unstable at body temperature with approximately $\frac{1}{4}$ of the activity of the *Val* polypeptide [4]. Thus, the *Met* allele is associated with lower enzyme activity, elevated dopamine levels and superior cognitive performance, while the *Val* allele is associated with higher enzyme activity, reduced dopamine levels and inferior cognitive performance [5, 6]. While feasible, the association of *COMT* and cognitive capacity has been the focus of debate. Some studies identified an association, such that individuals with a homozygous *Met/Met* genotype exhibit increased efficiency and superior performance on tests of executive functions and working memory in comparison to individuals with a homozygous *Val/Val* genotype [4, 7, 8]. By contrast, other studies, including a meta-analysis [9], describe this association to be rather limited or even absent [for a review, see 5; 8, 10].

The respective discrepancies may partially arise from differential effects of the *COMT* polymorphism on cognitive capacity during midlife development and old age. It has been put forward that the impact of genetic polymorphisms on cognitive capacity increases when resources decline – as is the case in aging [11]. As such, a decline in anatomical and neurochemical brain resources may lead to subsequent decline in compensatory skills, thereby amplifying genetic effects on cognitive capacity. In line with this are results reported by Nagel et al. [12] who found a negative effect of the *Val* allele on the number of perseverative errors in the Wisconsin Card Sorting Test in older (between 60 and 70 years of age), but not younger (between 20 and 30 years of age), participants. However, longitudinal studies yield contradictory results. De Frias et al. [13] found *Val/Val* carriers' performance on tasks of executive functioning to decline over a 5-year interval in contrast to *Met* carriers and identified a *COMT* × age interaction for middle-aged adults (50–60 years). Fiocco et al. [14] found the opposite, such that individuals (aged 70–79) with a homozygous *Val/Val* genotype displayed significantly less decline in the performance on the Digit Symbol Substitution Test than carriers of the *Met/Met* genotype across an 8-year interval, indicating an approximation of genotypes in cognitive test performance in aging. Other longitudinal studies have found no genetic impact on cognitive decline (n = 53, mean age 75.5, SD = 5.3 [15]; n = 473, age 64–68 [16]).

No overall conclusion can be drawn based on the respective studies, as they differ regarding their overall design, neuropsychological instruments, age of subjects, and length of follow-up interval. Previous research has focused on cognitive domains most commonly linked to PFC activity, e.g. executive functioning and working memory [4, 17], while from a neuropsychological standpoint, the inclusion of other cognitive domains relying on PFC activation (e.g. episodic memory retrieval [18]) is important to investigate the domain specificity of the respective effects. Here, we sought to examine the role of the *COMT p.Val158Met* genotype in different aspects of cognitive capacity in the course of adult development, healthy aging and the development of mild cognitive impairment (MCI) in the Interdisciplinary Longitudinal Study on Adult Development and Aging (ILSE), which involves two large birth cohorts of subjects born between 1930 and 1932 (C30) or between 1950 and 1952 (C50) [19, 20]. We hypothesized that there are specific effects of *COMT p.Val158Met* polymorphism on tests of executive functioning, which are more pronounced in older than in younger subjects.

Methods

Participants

Participants were recruited via local registries. For the purpose of the present study, exclusion criteria were psychiatric diagnoses affecting cognitive functioning apart from MCI, as defined by the Aging-Associated Cognitive Decline criteria [21]. Participants with mild cognitive disorder due to a medical condition, manifest Alzheimer's disease, other forms of dementia, or mood disorders were excluded. Examinations of both birth cohorts were conducted in parallel.

Measures

The first examination (T1) took place in 1993/1994, the second examination (T2) in 1998/1999, and the third examination (T3) was conducted between 2006 and 2008. Each time, careful screening of physical and mental health using extensive physical examination and the German version of the Structured Clinical Interview for the Diagnostic and Statistical Manual of Mental Disorders [22] was performed by trained physicians. DNA was extracted at T3 from whole blood using the Nucleon® Genomic DNA Extraction Kit BACC1. The *COMT* genotype was determined as a restriction fragment length polymorphism after PCR amplification and digestion with *Nla*III, as described by Lachman et al. [23]. The same kit was used for both cohorts.

To assess cognitive capacity, the subtests Word List (WL) and Digit Symbol Test (DST) of the Nuremberg Age Inventory [24], the subtests Mosaic Test (MT) and Finding Similarities (FS) of the Wechsler Intelligence Test Battery [25], the subtests Word Fluency (WF) and Visual Thinking (VT) of the Performance Evaluation System [26] as well as the Attention Endurance Test 'd2' (D2)[27] were administered. Due to time restrictions, certain subtests were not administered to the younger birth cohort at T2.

Statistical Analyses

Statistical analyses were performed using the SPSS 14.0 statistical package. After data description, analyses of variance (ANOVAs) and χ^2 tests were conducted to test for significant differences between (a) cohorts, healthy participants and participants with MCI, and (b) carriers of different *COMT p.Val158Met* genotypes. Afterwards, repeated-measures ANOVAs were performed with test scores at all three examination waves being treated as repeated measures. A separate analysis was conducted for healthy individuals using cohort (C30/C50) and *COMT* genotypes as independent variables and controlling for the level of education. Afterwards, C30 was analyzed separately to allow for the inclusion of cognitive status (MCI) in the model. In case assumption of sphericity was violated, Greenhouse-Geisser corrected values were used. The Bonferroni correction was applied to correct for multiple testing.

Results

Demographics and Baseline Characteristics

A total of 587 participants were included in the analysis: 188 healthy individuals from C30, 93 individuals diagnosed with MCI from C30, and 306 healthy individuals from C50, respectively. Distribution of genotypes was consistent with the Hardy-Weinberg equilibrium (*Val/Val* = 21.98%, *Val/Met* = 52.30%; *Met/Met* = 25.72%; $\chi^2 = 1.32$; $p = 0.251$). Demographic and baseline characteristics across genotypes can be inferred from table 1, while demographic characteristics across cohorts can be inferred from table 2.

Cognition

Results of the repeated-measures ANOVAs are presented in table 3. For the healthy participants from C30 and C50, no significant main effect of the *COMT* genotype on cognitive performance was identified. Cohort effects were evident for subtests DST (mean difference = -6.897; SE = 0.847; $p < 0.001$), MT (mean difference = -2.467; SE = 0.716; $p = 0.001$), VT (mean difference = -2.103; SE = 0.532; $p < 0.001$), WL (mean difference = -1.783; SE = 0.236; $p < 0.001$), D2 (mean difference = -14.888; SE = 7.023; $p = 0.035$) and FS (mean difference = -0.814; SE = 0.323; $p = 0.012$), with C50 performing better than C30. A significant inter-

Table 1. Demographic characteristics and baseline cognitive performance across the *COMT* genotype

	Val/Val (n = 129)	Val/Met (n = 307)	Met/Met (n = 151)	ANOVA/ χ^2	Duncan
<i>Demographics</i>					
Age baseline	52.98 (9.30)	52.78 (9.43)	53.72 (9.28)	F = 0.521, p = 0.594	
Education	13.71 (2.74)	13.94 (2.77)	13.95 (2.74)	F = 0.382, p = 0.683	
APOE genotype (% $\epsilon 4$ allele)	19.38	25.41	18.54	$\chi^2 = 3.39$, d.f. = 2 p = 0.183	
Sex (% females)	48.83	48.86	47.68	$\chi^2 = 0.062$, d.f. = 2 p = 0.970	
Cohort (% C30)	47.29	46.58	50.99	$\chi^2 = 0.813$, d.f. = 2 p = 0.666	
Cognitive status (% MCI)	15.50	15.64	16.55	$\chi^2 = 0.079$, d.f. = 2 p = 0.961	
<i>Cognitive performance baseline</i>					
DST	48.29 (10.95)	51.10 (10.31)	50.61 (10.02)	F = 3.40, p = 0.034	Val/Val < Val/Met, Met/Met
MT	29.77 (8.07)	30.68 (8.44)	30.73 (8.13)	F = 0.632, p = 0.532	
WL	12.70 (3.16)	12.76 (3.40)	12.81 (3.35)	F = 0.038, p = 0.963	
D2	147.51 (42.52)	156.21 (33.12)	159.19 (38.85)	F = 3.77, p = 0.023	Val/Val < Val/Met, Met/Met
FS	26.02 (4.33)	26.70 (4.17)	26.43 (4.18)	F = 1.204, p = 0.301	
WF	31.26 (7.98)	33.11 (9.00)	32.40 (9.28)	F = 2.008, p = 0.135	
VT	23.33 (6.69)	24.27 (6.19)	24.38 (5.96)	F = 1.242, p = 0.290	

Figures in parentheses are SD.

Table 2. Demographic characteristics and baseline cognitive performance across cohorts and cognitive status groups

	Healthy		MCI C30 (n = 93)	ANOVA/ χ^2	Duncan
	C30 (n = 188)	C50 (n = 306)			
<i>Demographics</i>					
Mean age baseline \pm SD, years	62.78 \pm 0.897	44.15 \pm 0.904	62.76 \pm 0.877	F = 31525.63, p < 0.001	Healthy C50 < Healthy C30, MCI
Mean education \pm SD, years	13.77 \pm 3.027	14.47 \pm 2.525	12.25 \pm 2.170	F = 25.48, p < 0.001	Healthy C50 > Healthy C30 > MCI
APOE genotype (% $\epsilon 4$ allele)	22.34	22.55	21.51	$\chi^2 = 0.03$, d.f. = 2, p = 0.983	
Sex (% females)	53.19	46.08	47.31	$\chi^2 = 2.427$, d.f. = 2, p = 0.297	

action of *COMT* and cohort arose for DST, suggesting that an effect of the *COMT* genotype on executive functioning was apparent in C50 but not in C30 (fig. 1). A triple interaction of the *COMT* genotype, time and cohort was found for MT as illustrated in figure 2.

The second repeated-measures analysis examined the influence of the *COMT* genotype and diagnosis of MCI on cognitive performance over 14 years in C30. No significant effects of the *COMT* genotype appeared. Significant main effects of diagnosis (MCI/cognitively healthy) emerged for DST (mean difference = 8.018; SE = 2.252; p < 0.001), MT (mean difference = 3.693; SE = 0.934; p < 0.001), VT (mean difference = 2.689; SE = 0.770; p = 0.001), WF (mean difference = 4.902; SE = 1.091; p < 0.001), WL (mean difference = 1.427; SE = 0.325; p < 0.001),

Table 3. Results of the repeated-measures analyses

Test Healthy subjects C30 and C50	C30 (healthy and MCI)
<p>DST COMT: $F_{2,451} = 1.149, p = 0.318$ Cohort: $F_{1,451} = 66.263, p < 0.001$ COMT \times cohort: $F_{2,451} = 3.326, p = 0.037$ Time: $F_{2,902} = 0.258, p = 0.772$ COMT \times time: $F_{4,902} = 0.152, p = 0.962$ Time \times cohort: $F_{2,902} = 9.765, p < 0.001$ Time \times COMT \times cohort: $F_{4,902} = 0.522, p = 0.719$</p>	<p>COMT: $F_{2,249} = 0.972, p = 0.380$ Diagnosis: $F_{1,249} = 4.992, p < 0.001$ COMT \times diagnosis: $F_{2,249} = 1.8, p = 0.167$ Time: $F_{2,498} = 0.653, p = 0.521$ COMT \times time: $F_{4,498} = 0.701, p = 0.591$ Time \times diagnosis: $F_{2,498} = 0.951, p = 0.387$ Time \times COMT \times diagnosis: $F_{4,498} = 1.453, p = 0.215$</p>
<p>MT COMT: $F_{2,454} = 0.288, p = 0.750$ Cohort: $F_{1,454} = 11.885, p = 0.001$ COMT \times cohort: $F_{2,454} = 0.871, p = 0.419$ Time: $F_{1,454} = 0.053, p = 0.817$ COMT \times time: $F_{2,454} = 1.162, p = 0.314$ Time \times cohort: $F_{1,454} = 8.350, p = 0.004$ Time \times COMT \times cohort: $F_{2,454} = 4.909, p = 0.008$</p>	<p>COMT: $F_{2,250} = 0.622, p = 0.538$ Diagnosis: $F_{1,250} = 14.156, p < 0.001$ COMT \times diagnosis: $F_{2,250} = 1.113, p = 0.330$ Time: $F_{2,500} = 3.561, p = 0.029$ COMT \times time: $F_{4,500} = 2.127, p = 0.076$ Time \times diagnosis: $F_{2,500} = 0.267, p = 0.766$ Time \times COMT \times diagnosis: $F_{4,500} = 0.572, p = 0.683$</p>
<p>VT COMT: $F_{2,453} = 1.820, p = 0.163$ Cohort: $F_{1,453} = 15.644, p < 0.001$ COMT \times cohort: $F_{2,453} = 0.143, p = 0.867$ Time: $F_{1,453} = 2.170, p = 0.141$ COMT \times time: $F_{2,453} = 0.265, p = 0.767$ Time \times cohort: $F_{1,453} = 20.291, p < 0.001$ Time \times COMT \times cohort: $F_{2,453} = 0.536, p = 0.585$</p>	<p>COMT: $F_{2,249} = 1.572, p = 0.210$ Diagnosis: $F_{1,249} = 15.193, p < 0.001$ COMT \times diagnosis: $F_{2,249} = 0.277, p = 0.758$ Time: $F_{2,498} = 3.692, p = 0.026$ COMT \times time: $F_{4,498} = 1.976, p = 0.097$ Time \times diagnosis: $F_{2,498} = 2.942, p = 0.054$ Time \times COMT \times diagnosis: $F_{4,498} = 0.681, p = 0.605$</p>
<p>WF COMT: $F_{2,453} = 1.247, p = 0.288$ Cohort: $F_{1,453} = 0.003, p = 0.960$ COMT \times cohort: $F_{2,453} = 0.367, p = 0.693$ Time: $F_{1,453} = 3.352, p = 0.068$ COMT \times time: $F_{2,453} = 1.199, p = 0.302$ Time \times cohort: $F_{1,453} = 9.596, p = 0.002$ Time \times COMT \times cohort: $F_{2,453} = 2.766, p = 0.064$</p>	<p>COMT: $F_{2,248} = 0.334, p = 0.717$ Diagnosis: $F_{1,248} = 24.633, p < 0.001$ COMT \times diagnosis: $F_{2,248} = 0.264, p = 0.768$ Time: $F_{2,496} = 0.410, p = 0.664$ COMT \times time: $F_{4,496} = 0.971, p = 0.423$ Time \times diagnosis: $F_{2,496} = 7.027, p = 0.001$ Time \times COMT \times diagnosis: $F_{4,496} = 1.188, p = 0.315$</p>
<p>WL COMT: $F_{2,451} = 0.622, p = 0.537$ Cohort: $F_{1,451} = 57.018, p < 0.001$ COMT \times cohort: $F_{2,451} = 1.614, p = 0.200$ Time: $F_{2,902} = 0.225, p = 0.798$ COMT \times time: $F_{4,902} = 0.602, p = 0.661$ Time \times cohort: $F_{2,902} = 0.546, p = 0.580$ Time \times COMT \times cohort: $F_{4,902} = 0.616, p = 0.651$</p>	<p>COMT: $F_{2,250} = 0.689, p = 0.503$ Diagnosis: $F_{1,250} = 26.705, p < 0.001$ COMT \times diagnosis: $F_{2,250} = 0.095, p = 0.909$ Time: $F_{2,500} = 9.960, p < 0.001$ COMT \times time: $F_{4,500} = 0.099, p = 0.983$ Time \times diagnosis: $F_{2,500} = 3.969, p = 0.019$ Time \times COMT \times diagnosis: $F_{4,500} = 1.865, p = 0.115$</p>
<p>D2 COMT: $F_{2,442} = 1.237, p = 0.291$ Cohort: $F_{1,442} = 4.494, p = 0.035$ COMT \times cohort: $F_{2,442} = 1.269, p = 0.282$ Time: $F_{2,884} = 12.590, p < 0.001$ COMT \times time: $F_{4,884} = 0.152, p = 0.962$ Time \times cohort: $F_{2,884} = 31.965, p < 0.001$ Time \times COMT \times cohort: $F_{4,884} = 0.823, p = 0.511$</p>	<p>COMT: $F_{2,241} = 0.355, p = 0.702$ Diagnosis: $F_{1,241} = 32.309, p < 0.001$ COMT \times diagnosis: $F_{2,241} = 0.802, p = 0.450$ Time: $F_{2,482} = 9.122, p < 0.001$ COMT \times time: $F_{4,482} = 0.992, p = 0.412$ Time \times diagnosis: $F_{2,482} = 1.143, p = 0.320$ Time \times COMT \times diagnosis: $F_{4,482} = 0.556, p = 0.694$</p>
<p>FS COMT: $F_{2,454} = 0.013, p = 0.987$ Cohort: $F_{1,454} = 6.353, p = 0.012$ COMT \times cohort: $F_{2,454} = 0.09, p = 0.914$ Time: $F_{1,454} = 0.140, p = 0.709$ COMT \times time: $F_{2,454} = 1.561, p = 0.211$ Time \times cohort: $F_{1,454} = 1.899, p = 0.169$ Time \times COMT \times cohort: $F_{2,454} = 0.339, p = 0.713$</p>	<p>COMT: $F_{2,251} = 1.147, p = 0.319$ Diagnosis: $F_{1,251} = 31.246, p < 0.001$ COMT \times diagnosis: $F_{2,251} = 0.298, p = 0.743$ Time: $F_{2,502} = 4.832, p = 0.008$ COMT \times time: $F_{4,502} = 2.352, p = 0.053$ Time \times diagnosis: $F_{2,502} = 3.814, p = 0.023$ Time \times COMT \times diagnosis: $F_{4,502} = 1.784, p = 0.131$</p>

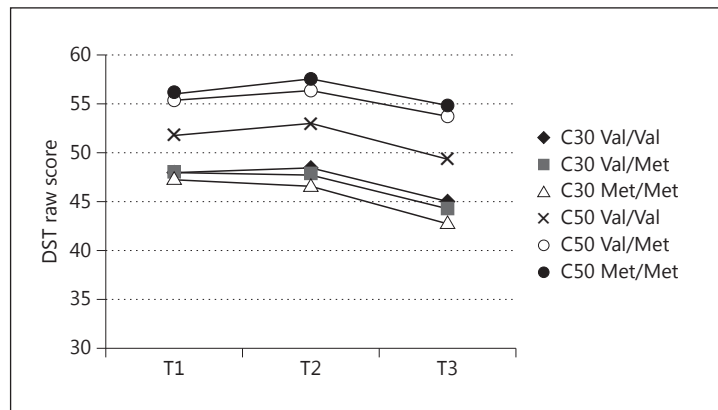


Fig. 1. Interaction of *COMT* and cohort for DST.

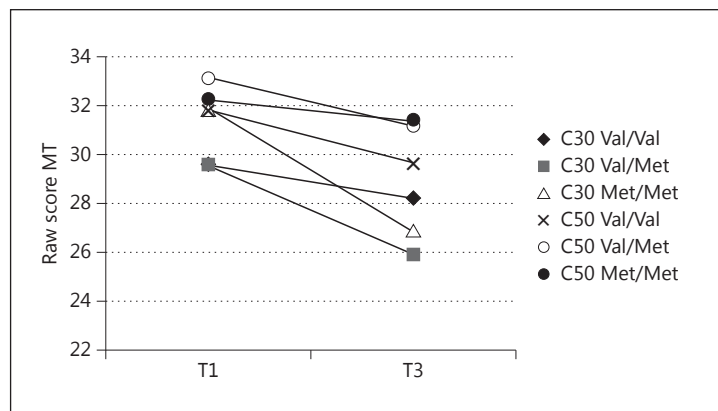


Fig. 2. Interaction of *COMT*, cohort and time for MT.

D2 (mean difference = 57.102; SE = 10.046; $p < 0.001$), and FS (mean difference = 2.920; SE = 0.515; $p < 0.001$), with healthy individuals outperforming those diagnosed with MCI. No interaction of the *COMT* genotype and diagnosis was found. No significant interaction effects of the *COMT* genotype and time appeared, while interaction effects by trend emerged for MT, VT, and FS. Significant interaction effects of time and diagnosis of MCI were found for WF, WL, and FS. No triple interaction was observed.

Discussion

In this study, we investigated the effect of the *COMT p.Val158Met* polymorphism on cognitive performance in a sample of 587 participants of two distinct age cohorts, born between 1930 and 1932 (C30) or between 1950 and 1952 (C50). Our results suggest that the *COMT* genotype exerts a different influence on cognitive functioning for the C50 than for the C30 cohort. For C50, we find significant differences in baseline test performance between *COMT* genotypes on the subtests D2 and DST, such that homozygous *Val* carriers perform more slowly than heterozygotes and homozygous *Met* carriers. We identified an interaction suggesting that this effect is only applicable to C50, but not to C30, contrary to our second hypothesis. The minimal effects of the *COMT p.Val158Met* polymorphism are more pronounced in tests of executive functioning than other cognitive domains. However, no interaction with time was identified, indicating that the *COMT* genotype does not influence cognitive trajectories over time. An individual analysis for the C30 cohort suggests that cognitive perfor-

mance trajectories in older subjects are largely independent on the *COMT* genotype. No effect was found for subjects diagnosed with MCI. These results confirm the importance of the *COMT p.Val158Met* genotype on tasks assessing attention and cognitive flexibility in midlife but not in healthy aging and the development of MCI.

De Frias et al. [13] found that the performance of *Val/Val* carriers on tasks of executive functioning declined over a 5-year interval compared to that of *Met* carriers. An interaction of *COMT* and age was identified for middle-aged participants (aged 50–60), supporting the idea that discrepancies due to genetic effects are greater in midlife than in aging. Fiocco et al. [14] identified a difference in cognitive decline across an 8-year interval, such that individuals with a homozygous *Val/Val* genotype displayed significantly less decline in the Digit Symbol Substitution Test performance compared to *Met* homozygotes, indicating an approximation of different genotypes in test performance in the course of healthy aging. Generally, the respective studies are in line with our findings, even though we did not identify an interaction effect of time (age) and the *COMT* genotype. However, a few studies point to a potential amplification of genetic effects in old age [12]. While it is plausible that losses of brain resources such as decline of striatal and extrastriatal dopamine or atrophy affecting the PFC may amplify the effects of genetic polymorphisms such as *COMT p.Val158Met* on cognition [11], our results show that the *COMT* genotype on its own is not a determining factor. Further studies have demonstrated inefficient cortical processing as reflected by low performance and greater activity in *Val* homozygotes compared to *Met* homozygotes in tasks demanding working memory capacity in participants in their mid-thirties [4, 28] and attentional control [29]. Remarkably, neurological differences were sometimes identified in the absence of effects on behavioral measures such as test performance [28], suggesting a compensatory mechanism. Since we did not find an effect of the *COMT* genotype on cognitive trajectories, we must consider that certain factors related to the birth cohort are determinative rather than age per se.

A potential limitation to studies examining specific cognitive domains is their reliance on neuropsychological test batteries that are largely classified by their content. Assessment instruments can only partially reflect differential cognitive domains or phenotypes (for a review, see Harris and Deary [30]). Moreover, an interplay of different candidate genes affecting dopamine regulation seems likely. There exists relatively robust evidence for risk of increased cognitive decline from *APOE ε4* allele as well as *BDNF* [30, 12]. However, in this study, we were able to consider a follow-up interval of 14 years, allowing for conclusions on the influence of the *COMT* genotype on the process of healthy aging and the development of MCI, while previous research was limited to a few years only. Moreover, directly contrasting two different birth cohorts allowed us to delineate cohort effects from aging effects. Our findings can shed light on the often somewhat contradictory findings reported in the literature. Another strength of this study is the use of extensive neuropsychological testing. Given the role of the *COMT* genotype in dopaminergic pathways, it is likely that areas relying on the PFC are affected differently than other areas. Results of our study suggest that the *COMT p.Val158Met* polymorphism has a larger genetic contribution to tests of attention, cognitive flexibility and information processing speed at ages 43–56 than at ages 63–76. The effects of *COMT* were therefore specific to tests assessing executive functioning rather than tests of memory, verbal fluency or visuospatial thinking.

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